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# Genome-wide association and replication studies identify four variants associated with prostate cancer susceptibility 

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#### Abstract

We report a genome-wide association follow up study on prostate cancer. We identify four variants associated with the disease in European populations: rs $10934853-\mathrm{A}(\mathrm{OR}=1.12, \mathrm{P}=$ $2.9 \times 10^{-10}$ ) on 3 q 21.3 , two moderately correlated ( $\mathrm{r}^{2}=0.07$ ) variants on 8 q 24.21 ; rs16902094-G $\left(\mathrm{OR}=1.21, \mathrm{P}=6.2 \times 10^{-15}\right)$ and rs445114-T $\left(\mathrm{OR}=1.14, \mathrm{P}=4.7 \times 10^{-10}\right)$ and rs8102476-C $(\mathrm{OR}=$ $1.12, \mathrm{P}=1.6 \times 10^{-11}$ ) on 19 q 13.2 . We also refine a previous association signal on 11 q 13 with the SNP rs11228565-A ( $\mathrm{OR}=1.23, \mathrm{P}=6.7 \times 10^{-12}$ ). In a multi-variant analysis, using 22 prostate cancer risk variants typed in the Icelandic population, we estimate that carriers belonging to the top $1.3 \%$ of the risk distribution have a risk of developing the disease that is more than 2.5 times greater than the population average risk estimates.


We and others have previously presented results from genome-wide association studies (GWAS) on prostate cancer reporting several common variants conferring risk of the disease ${ }^{1-7}$. By scrutinizing both our Icelandic GWAS data and publicly available data, as well as through fine-mapping work of two previously published loci on 8q24.21 and 11q13, we identified four new variants conferring risk of prostate cancer and a variant refining the previously published association signal on 11q13.

In order to search for prostate cancer risk variants, we performed an analysis of the combined data from the Icelandic GWAS, generated using the Illumina 317 K chip, and from the replication genotyping project, released (spring 2008) by the National Cancer Institute Cancer GeneticMarkers of Susceptibility (CGEMS), using 25K SNPs genotyped on five study populations, including PLCO, CPS-II, HPFS, FPCC and ATBC ${ }^{5,7}$. In the combined study, when excluding regions containing previously reported prostate cancer risk variants, only two SNPs have $P<1 \times 10^{-5}$. These two new variants are: allele A of rs 10934853 (rs10934853-A) located on 3q21.3 (a fully correlated marker used from the CGEMS data is rs4857841 with D' and $\mathrm{r}^{2}=1$ according to CEU HapMap, Supplementary Table 1); allelic odds ratio (OR) of 1.14 and $P=1.6 \times 10^{-7}$ and allele C of rs8102476 (rs8102476-C) located on 19q13.2; OR of 1.11 and $P=2.6 \times 10^{-6}$. In the combined study of the Icelandic and CGEMS data, when analyzing variants located within 1 Mb regions of previously published prostate cancer risk loci, we found only the SNP rs 445114 , located on 8 q 24.21 , to be associated with prostate cancer $\left(\mathrm{OR}=1.14, P=3.1 \times 10^{-8}\right)$ and not to be correlated with other signals at any of the previously reported prostate cancer loci.

As a part of ongoing fine-mapping project on the three previously published variants on 8 q 24.21 we re-sequenced a 527 kb candidate region on 8 q 24 using pools of either Icelandic case or control samples (see Online Methods). We prioritized the analysis of 7 SNPs producing suggestive association results for prostate cancer in the analysis of the pooled samples (Supplementary Table 1). These 7 SNPs were genotyped, using Centaurus single track assay, in our Icelandic set of up to 1,980 patients and 7,000 controls. Six of the SNPs were found to be correlated with one of the previously reported variants on 8 q 24 , while rs16902094 was not found to be correlated with any of the previously reported prostate cancer variants on 8q24.21 (Supplementary Tables 1 and 2). Furthermore, rs16902094 was found to be associated with prostate cancer in Iceland, having an OR of $1.28\left(P=3.5 \times 10^{-6}\right.$; Table 1). rs16902094 is not present on the Illumina Hap 317K SNP chip and no data are
available for it in the HapMap project according to HapMap's official website. However, by genotyping SNPs located within the same LD-region as rs16902094 and that are present on the Illumina Hap 550 chip, we identified a highly correlated SNP rs16902104 ( $\mathrm{D}^{\prime}=0.98, \mathrm{r}^{2}$ $=0.96$ according to data from 2,633 Icelanders; Supplementary Table 1) which CGEMS has released data for. By combining the Icelandic data for rs 16902094 and the CGEMS data for rs16902104 we obtained an OR of 1.21 ( $\mathrm{P}=1.1 \times 10^{-9}$ ).

The two new SNPs on 8q24.21, rs16902094 and rs445114, are located in the same LDregion but the correlation between them is very low ( $D^{\prime}=1$ and $r^{2}=0.07$ according to data from 5,450 Icelanders; Supplementary Table 1) and the results for both remain significant after adjustment for the other (Supplementary Table 3a). This suggests that a unique variant capturing the effect of both rs16902094 and rs445114 remains to be discovered or, alternatively, that the LD-region contains more than one variant that predisposes to prostate cancer. Of the previously published cancer variants on 8 q 24 , only the breast cancer variant (rs13281615) $)^{8}$ is located within the same LD-region as the two new 8 q 24 SNPs and rs445114 is somewhat correlated with it ( $D^{\prime}=0.76, r^{2}=0.44$; Supplementary Table 2), while rs16902094 is much less correlated with it ( $D^{\prime}=0.61, r^{2}=0.06$; Supplementary Table 2). However, both rs 16902094 and rs 445114 show very little correlation with any of the previously published prostate ${ }^{1,2,7,9}-$, colon ${ }^{10-12}$-, or bladder cancer ${ }^{13}$ risk variants on 8 q 24 ( ${ }^{\prime} \leq 0.6$ and r ${ }^{2} \leq 0.13$; Supplementary Table 2 and Supplementary Fig. 1). The results in Iceland for rs16902094, rs445114 and the three previously published prostate cancer risk variants on 8 q24, remain significant after being adjusted for each other (Supplementary Table 3b). Hence, rs16902094 and rs445114 should be included in the list of variants at $8 q 24$ associated with prostate cancer risk. The five $8 q 24$ variants conferring risk of prostate cancer are distributed between four LD-regions, spanning approximately 480 kb . One of the five prostate cancer risk variants, rs6983267, has been shown to also affect the risk of colorectal cancer. Other cancer risk variants on 8 q24.21, conferring risk of breast or bladder cancer, have not been shown to predispose to prostate cancer ${ }^{13,14}$.

We proceeded to genotype all the four newly discovered SNPs (rs10934853 on 3q21.3, rs8102476 on 19q13.2 and rs16902094 and rs16902104 both on 8q24.21) in at least two out of five prostate cancer study groups (deCODE follow up groups) of European descent. These groups come from The Netherlands, Spain, Finland and the United States (US) (see Supplementary Note). When results for SNPs successfully genotyped in these groups were combined with the Icelandic and CGEMS data discussed above, they were significant for all loci, surpassing the threshold of genome-wide significance set by many at $P>10^{-7}$. By examining the literature we found that Duggan et al. ${ }^{15}$ had published data $(\mathrm{OR}=1.21, \mathrm{P}=$ 0.045 ) for rs10934853, on 3q21.3, from a study on aggressive prostate cancer in the CAPS study population from Sweden. Thereby, adding further significance to the combined analysis of the 3 q 21.3 locus (see Table 1 for combined OR and $P$-values).

When inspected, a test of heterogeneity in the OR for all variants and all study groups showed a nominally significant heterogeneity $\left(P_{\text {het }}=0.039\right)$ for the 3 q 21.3 locus, no significant difference was observed for the other three loci ( $P_{\text {het }}>0.1$ ). For rs10934853 on $3 q 21.3$ the estimated OR tended to be greater for the combined US study groups ( $\mathrm{OR}=1.21$ and the $95 \% \mathrm{CI}: 1.13,1.28$ ) than for the combined European study groups $(\mathrm{OR}=1.08$ and the $95 \% \mathrm{CI}: 1.04,1.13$ ). This observation, while interesting, needs to be further confirmed.

A prostate cancer risk variant, with an allelic frequency of about 0.5 and located on 11q13, was reported previously by two groups independently; Thomas et al. ${ }^{5}$ reported rs10896449, and Eeles et al. ${ }^{4}$ reported rs7931342 (the two SNPs are highly correlated with D' $=1$ and $\mathrm{r}^{2}$ $=0.97$ according to Utah CEPH (CEU) HapMap data). In the Icelandic GWAS data set, the strongest association with prostate cancer on 11q13 was observed for allele G of rs 10896450
( $\mathrm{OR}=1.13, P=2.5 \times 10^{-4}$; Table 2 ), a SNP highly correlated with the two previously reported SNPs $\left(D^{\prime}=1 ; r^{2}>0.97\right.$ according to CEU HapMap data). In order to investigate further the 11q13 locus, we selected six SNPs not present on the Illumina Hap300 chip but moderately correlated with each of the three anchor SNPs: rs10896449, rs7931342, and rs10896450 ( $\mathrm{D}^{\prime} \geq 0.70$ and $\mathrm{r}^{2}>0.24$ according to CEU HapMap data; Supplementary Table 1). Furthermore, these SNPs are not strongly correlated with any SNP on the Illumina Hap300 chip ( $r^{2}<0.8$ ). By doing this, we attempted to find variants with greater risk and lower allelic frequency. The additional six SNPs were genotyped in 1,809 and 783 Icelandic cases and controls, respectively. The control individuals were randomly selected from the Icelandic population and were not known to have prostate cancer according to the nationwide Icelandic Cancer Registry (see URL below). Allele A of the refinement SNP rs11228565 was found to confer greater risk of the disease than the anchor SNP (rs10896450), with an OR of $1.24(P=0.0057)$ and a control frequency of 0.177 in the Icelandic study group (Table 1 and Supplementary Table 4). We then tested the two SNPs, rs10896450 and rs11228565, in the five deCODE follow up groups. For the Finnish study group, instead of genotyping rs10896450, we used data available for the highly correlated SNP rs7931342 ( $\mathrm{D}^{\prime}=1 ; \mathrm{r}^{2}>0.97$ according to CEU HapMap data), previously reported by Kote-Jarai et al ${ }^{16}$. Combination of the results from Iceland and the deCODE follow up groups gave an estimated OR of $1.15\left(\mathrm{P}=2.6 \times 10^{-8}\right)$ for rs10896450-G whereas for rs1128565-A the OR was estimated to be $1.23\left(P=6.7 \times 10^{-12}\right)$ (Table 1 and Supplementary Table 5). The estimated frequency of the risk allele of rs11285565 in the combined data set of the groups used in this study is 0.20 compared to 0.50 for the G allele of rs 10896450 (Supplementary Table 5). Hence, the frequency of the risk allele for the new SNP is lower than for the previously reported one. Also, the point estimate of the OR is greater ( $\mathrm{OR}=$ 1.23 for rs11285565 vs. 1.15 for rs10896450) although the difference is not significant, ( $\mathrm{P}=$ 0.08 using results for all study groups; Supplementary Table 5). The results for rs 11228565 were strongest for the study groups from Nashville and Finland but weakest for the Spanish group. However, a test of heterogeneity in the OR of all study groups showed no significant difference for the 11 q 13 locus $(P=0.26)$. When adjusting the results for the refinement SNP (rs11228565-A) using data the anchor SNPs, (rs10896450-G or rs7931342-G), it remained significant $\left(\mathrm{OR}=1.19, \mathrm{P}=5.0 \times 10^{-7}\right)$ while the anchor SNPs were not significant after being adjusted for the refinement $\mathrm{SNP}(\mathrm{OR}=1.03, \mathrm{P}=0.15$; Supplementary Table 5).

For the six study groups of European descent (excluding the five CGEMS groups and the study by Duggan et al ${ }^{15}$.), where we had information about age at diagnosis, no age effect was seen for any of the five loci discussed above ( $P \geq 0.3$ ). By computing the genotype specific ORs or inspecting the genotypic ORs from the public CGEMS data set for the variants in Table 1, we found that the multiplicative model provides an adequate fit for all five loci in the study groups analyzed (for the full- vs. the multiplicative model all $\mathrm{P}>0.1$; Supplementary Table 6 and Online Methods). Comparing the patients with a more aggressive phenotype (Gleason $\geq 7$ and/or T3 or higher and/or node positive and/or metastatic disease) to the group with less aggressive tumors (Gleason <7 and T 2 or lower) showed no difference; an OR between 0.94 and $1.03(P>0.1)$ was observed for the five loci. The public CGEMS data also showed no difference between patients with more or less aggressive disease ( $P>0.40$ ) for the SNPs where data are available (3q21, 8q24, and 19q13.2). The study published by Duggan et al ${ }^{15}$. was on aggressive prostate cancer only and a comparison with less aggressive disease was therefore not possible.

The four novel loci reported here, add to the rapidly increasing number of prostate cancer susceptibility variants, identified through GWASs. In Table 2, we provide results from the Icelandic GWAS for risk variants reported by us and/or others (until February 2009) to confer risk of prostate cancer. For some of the SNP associations reported in other populations, the Icelandic results provide replication (in particular 7p15 and 10q11), while
we failed to replicate association for other SNPs in this population (in particular 9q33 and 22q13). In order to summarize the overall effect of the variants in Tables 1 and 2 we combined the effect of all variants affecting risk of prostate cancer in the Icelandic population., We performed a multi-variant analysis, using the multiplicative model for 22 risk variants. Testing the assumptions, no significant deviation from the multiplicative model was observed for any given variant as well as no interaction between variants (see Online Methods). Based on this analysis, the estimated risk is more than 2.5 -fold greater for the top $1.3 \%$ of the risk distribution (including $0.3 \%$ of the population with a risk greater than 3 ), using the population average risk as a reference (Table 3). For these individuals this corresponds to a lifetime risk of over $30 \%$ of being diagnosed with prostate cancer, compared with a risk of $12 \%$, on average in Iceland, of getting the disease before age 75, according to NORDCAN (see URL below) assuming no interaction between the effect of the variants and age at diagnosis. The combined risk estimates presented here are similar to those previously reported by Kote-Jarai et al. for 15 risk variants ${ }^{16}$. We note that the estimates provided here are based solely on our Icelandic study population, and that more accurate estimates could be obtained from large prospective studies. Also, given the fast pace of discoveries in the current era of GWASs, more variants associated with prostate cancer risk are likely to be discovered, which suggests the need for constant updating of such multivariant risk models.

## Online Methods

## Genotyping Methods

Illumina genotyping-1,968 and 35,382 Icelandic case- and control-samples respectively, were successfully assayed with the Infinium HumanHap300 SNP chip (Illumina, SanDiego, CA, USA), containing 317,503 haplotype tagging SNPs derived from phase I of the International HapMap project. Of the SNPs assayed on the chip, 2,906 SNPs had a yield lower than $95 \%, 271$ SNPs had a minor allele frequency, in the combined set of cases and controls, below 0.01 or were monomorphic. An additional 4,632 SNPs showed a significant distortion from Hardy-Weinberg equilibrium in the controls ( $P<1.0 \times 10^{-3}$ ). In total, 6,983 unique SNPs were removed from the study. Thus, the analysis reported in the main text utilizes 310,520 SNPs. Any samples with a call rate below $98 \%$ were excluded from the analysis.

Replication genotyping-Single SNP genotyping of the SNPs reported in the main text for the four case-control groups from Iceland, The Netherlands, Spain and Chicago was carried out by deCODE Genetics in Reykjavik, Iceland, applying the Centaurus ${ }^{20}$ (Nanogen) platform. The quality of each Centaurus SNP assay was evaluated by genotyping each assay in the CEU and/or YRI HapMap samples and comparing the results with the HapMap publicly released data. Assays with $>1.5 \%$ mismatch rate were not used and a linkage disequilibrium (LD) test was used for markers known to be in LD. We re-genotyped more than $10 \%$ of the samples and observed a mismatch rate lower than $0.5 \%$. Genotyping of samples from Finland and Nashville was done using the same Centaurus assays as used in Iceland at the University of Tampere and Vanderbilt University, respectively, using standard protocols.

For each of the SNPs discussed in the main text, the yield was higher than $95 \%$ for those samples which genotyping was attempted for in every study group. The SNPs rs 16902094 on $8 q 24$ and rs 11228565 on 11q13 are not present on the Human Hap300 chip. Therefore, using a single SNP assay for genotyping, an attempt was made to genotype 6,900 and 800 individuals, respectively, of the 35,382 Icelandic controls as well as 1,860 Icelandic cases and all available individuals from the replication study groups.

Discovery of new SNP on 8q24 by Solexa re-sequencing-In order to search for new SNPs on 8q24, a 527 kb region (128113108-128640337 bp, Build 36) was sequenced using the Solexa re-sequencing platform (Illumina Inc.). From our set of $\sim 2,000$ cases, 800 were selected randomly and split into two DNA-pools, each with 400 samples. Similarly, 800 control individuals, not known to have prostate cancer, were selected randomly and split into two DNA-pools. Dilutions were prepared in duplicates for the 4 pools and used for longe-range PCR reactions (each amplimer consisted of about 10 kb ). PCR fragments were run on $0.8 \%$ agarose gels and the DNA visualized with BlueView (Sigma Inc.) under normal light and their sizes estimated with HindIII digested lambda size marker (Fermentas Inc). Bands of correct sizes were excised out of the agarose gels and purified with Qiagen gel extraction kit (Qiagen Inc.). The PCR products were quantified by picogreen assay (Invitrogen Inc.) as described by the manufacturer. The preparation of the Solexa DNA libraries, the cluster generation and DNA sequencing was done was done as described by Bentley et $\mathrm{al}^{21}$. The SNP Analysis pipeline is composed of four components: Alignment, SNP calling, Filtering and Association analysis. Promising SNPs were selected for further study/confirmation using Centaurus single track SNP assays.

## Statistical analysis

Association analysis-For SNPs that were in strong LD, whenever the genotype of one SNP was missing for an individual, the genotype of the correlated SNP was used to provide partial information through a likelihood approach as previously described ${ }^{9}$. This ensured that results presented in Supplementary Table 5 were based on the same number of individuals, allowing meaningful comparisons of results for correlated SNPs. A likelihood procedure described in a previous publication ${ }^{22}$ and implemented in the NEMO software was used for the association analyses.

We tested the association of an allele to prostate cancer using a standard likelihood ratio statistic that, if the subjects were unrelated, would have asymptotically a $\chi^{2}$ distribution with one degree of freedom under the null hypothesis. Allelic frequencies rather than carrier frequencies are presented for the markers in the main text. Allele-specific ORs and associated $P$ values were calculated assuming a multiplicative model for the two chromosomes of an individual ${ }^{23}$. Results from multiple case-control groups were combined using a Mantel-Haenszel model ${ }^{24}$ in which the groups were allowed to have different population frequencies for alleles, haplotypes and genotypes but were assumed to have common relative risks (see Gudmundsson et al. ${ }^{3}$ for a more detailed description of the association analysis).

The control groups from Iceland, The Netherlands, Spain, and Finland include both male and female controls. No significant difference between male and female controls was detected for SNPs presented in Table 1 for each of these four groups. Controls from other study groups include only males.

In order to assess the association for the SNP rs4962416 on 10q26, which is in the CEU section of the Hapmap database but absent from the Illumina Hap300 chip, we use a method based on haplotypes of two markers (rs7077275 and rs893856) present on the chip. We used a method we have previously employed ${ }^{25}$, that is an extension of the two-marker haplotype tagging method ${ }^{26}$ and is similar in spirit to two other proposed methods ${ }^{27,28}$. We computed associations with a linear combination of the different haplotypes chosen to act as surrogates to HapMap markers in the regions. These calculations were based on 1,724 prostate cancer cases and 35,322 controls genotyped on chip.

Multivariant analysis-In a multi-variant analysis, we combined the effect of 22 variants affecting risk of prostate cancer using estimates based on data from only the Icelandic study
group. A multiplicative model was assumed at each variant and between all variants. For the 21 autosomal variants we tested for deviation from the multiplicative model comparing it to the full model of genotypic odds ratio (OR). No significant (given the number of tests performed) deviation from the multiplicative model was detected (all P-values $>0.0024$, corresponding to 0.05 divided by 21 ). For the variant located on the X chromosome (rs5945572), there exist only two male genotypes (carrier and no-carrier). Also, we tested the pair-wise interactions between the 22 risk variants using logistic regressions including terms corresponding to the 231 possible pairs. No significant (given the number of tests performed) deviation from the multiplicative model was detected with a level of significance below 0.00022 (corresponding to $\mathrm{P}=0.05$ divided by 231 ). Similarly, the absence of interaction between variants has previously been reported by others ${ }^{5,16,29}$. Odds ratios were calculated for all possible genotype combinations based on these 22 variants and expressed relative to the average general population risk. The combined OR estimates were then divided into OR-ranges and presented along with the percentage of the population within each OR-range (See Table 3 in the main text). The general population risk was determined using a frequency-weighted average risk for all possible genotypes.

The Icelandic samples were part of the initial discovery study populations for 10 of these 22 variants, therefore the estimates for these may be inflated due to winners curse. From Table 1 we are using the estimates for the following variants: rs10934853 on 3q21.3, rs11228565 on 11q13 and rs8102476 on 19q13.2. For the five variants (rs1447295, rs6983267, rs16901979, rs16902094 and rs445114) on 8q24.21 we are using the adjusted estimates as reported in Supplementary Table 3b. From Table 2, we use the estimates for all variants except rs 1571801 on 9 q 33 since its effect was in the opposite direction compared to the original publication, and rs10896450 on 11q13 for which data for the refinement SNP (rs11228565) in Table 1 was used. Furthermore, for the two variants on 17q12, reported in Table 2, we are using the estimates adjusted for each other since the two markers are located $\sim 25 \mathrm{~kb}$ to each other although not correlated ( $\mathrm{D}^{\prime}=0.03, \mathrm{r}^{2}=0.0004$ according to CEU HapMap data); rs4430796 has an adjusted OR of 1.17 and rs11649743 has an adjusted OR of 1.06 .

We apply the combined genetic relative risk to the lifetime risk by multiplying them together. Since the association with the disease for these variants has not been shown to depend on age at diagnosis ${ }^{3,16}$, we assume that the individual estimates for each variant will have a similar effect at any given age. The lifetime risk of getting prostate cancer is estimated to be $12 \%$, on average in Iceland, before the age of 75 , according to NORDCAN (see URL).

Analysis of the CGEMS data-For the five individual study populations from the CGEMS study ${ }^{5,7}$ (ACS = American Cancer Society Prevention Study II (US); ATBC = Alpha-Tocopherol, Beta-Carotene Prevention Study (Finland); FPCC: CeRePP French Prostate Case-Control Study (France); HPFS = Health Professionals Follow-up Study (US); PLCO = Prostate, Lung, Colon, Ovarian Trial (US)), when assessing the allelic effect we used the pre-computed data (released in spring, 2008) corresponding to "All case versus control (dichotomous), genotype trend effect model, adjusted". When assessing the genotypic effect at each loci for the CGEMS study we used the pre-computed "All case versus control (dichotomous), genotype-specific effect model, adjusted, ALL (ACS, HPFS, FPCC, ATBC, PLCO)".

Correction for relatedness-Some individuals in the Icelandic case-control groups were related to each other, causing the aforementioned $\chi^{2}$ test statistic to have a mean $>1$. We estimated the inflation factor by using a previously described procedure ${ }^{30}$ in which we simulated genotypes through the genealogy of the 37,350 Icelanders analyzed in the present
study (number of simulations $=100,000$ ). The inflation factor was estimated to be 1.10 . Results from the Icelandic samples presented in the main text are based on adjusting the $\chi^{2}$ statistics by dividing each of them by 1.10 .

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Summary association results for the SNPs on $3 q$ Table 1.3 , 8q24, 19q13.2 and 11

| Study population | $\begin{aligned} & \text { Cases } \\ & (\mathbf{N}) \end{aligned}$ | Controls <br> (N) | Frequency |  | OR (95\% CI) | P-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Cases | Controls |  |  |
| A. Results for rs10934853 [A] or rs4857841 [A] on 3q21.3 |  |  |  |  |  |  |
| Iceland ${ }^{\text {a }}$ | 1,968 | 35,227 | 0.295 | 0.269 | 1.14 (1.06, 1.22) | 3.2E-04 |
| Chicago, Illinois | 1,077 | 1,003 | 0.313 | 0.273 | 1.21 (1.06, 1.39) | 4.4E-03 |
| Finland | 2,638 | 1,716 | 0.330 | 0.319 | 1.05 (0.96, 1.15) | 0.27 |
| The Netherlands | 1,084 | 1,827 | 0.306 | 0.286 | 1.10 (0.98, 1.24) | 0.10 |
| Nashville, Tennessee | 596 | 687 | 0.283 | 0.270 | 1.07 (0.90, 1.27) | 0.47 |
| Spain | 811 | 1,605 | 0.306 | 0.314 | 0.96 (0.84, 1.09) | 0.54 |
| $\mathrm{ACS}^{b}$ | 1,758 | 1,775 | 0.300 | 0.258 | 1.25 (1.12, 1.39) | 4.3E-05 |
| $\text { ATBC }^{b}$ | 928 | 921 | 0.309 | 0.319 | 0.96 (0.84, 1.10) | 0.59 |
| $\operatorname{FPCC}^{b}$ | 654 | 657 | 0.291 | 0.272 | 1.09 (0.92, 1.29) | 0.34 |
| $\operatorname{HPFS}^{b}$ | 595 | 609 | 0.313 | 0.278 | 1.18 (0.99, 1.40) | 0.070 |
| $\mathrm{PLCO}^{b}$ | 1,167 | 1,093 | 0.308 | 0.266 | 1.23 (1.08, 1.41) | $2.5 \mathrm{E}-03$ |
| CAPS ${ }^{\text {c }}$ | 498 | 494 | 0.329 | 0.288 | 1.21 (1.00, 1.46) | 0.045 |


| All combined ${ }^{\text {d }}$ | 13,774 | 47,614 | - | 0.284 | 1.12 (1.08, 1.16) | 2.9E-10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathbf{P}_{\text {het }}$ |  |  |  |  | 0.039 |  |
| $\mathrm{I}^{2}$ |  |  |  |  | 46.4 |  |
| B. Results for rs16902094 [G] or rs16902104 [T] on 8q24 |  |  |  |  |  |  |
| Iceland ${ }^{\text {a }}$ | 1,858 | 6,853 | 0.168 | 0.136 | 1.28 (1.15, 1.41) | 3.5E-06 |
| Chicago, Illinois | 797 | 758 | 0.166 | 0.147 | 1.16 (0.95, 1.41) | 0.14 |
| Finland | 2,197 | 1,725 | 0.248 | 0.222 | 1.15 (1.03, 1.28) | $9.9 \mathrm{E}-03$ |
| The Netherlands | 831 | 837 | 0.161 | 0.138 | 1.20 (0.99, 1.44) | 0.066 |
| Nashville, Tennessee | 669 | 733 | 0.170 | 0.130 | 1.37 (1.11, 1.69) | 3.0E-03 |
| Spain | 643 | 952 | 0.162 | 0.137 | 1.21 (1.00, 1.48) | 0.055 |
| $\mathrm{ACS}^{\text {b }}$ | 1,759 | 1,774 | 0.156 | 0.132 | 1.22 (1.06, 1.39) | 4.3E-03 |


| Study population | $\begin{gathered} \text { Cases } \\ \text { (N) } \end{gathered}$ | $\underset{(\mathbf{N})}{\text { Controls }}$ | Frequency |  | OR (95\% CI) | P-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Cases | Controls |  |  |
| ATBC $^{b}$ | 929 | 920 | 0.255 | 0.193 | 1.43 (1.22, 1.67) | 1.0E-05 |
| FPCC ${ }^{\text {b }}$ | 656 | 657 | 0.152 | 0.145 | 1.06 (0.85, 1.31) | 0.61 |
| HPFs ${ }^{\text {b }}$ | 596 | 611 | 0.127 | 0.133 | 0.93 (0.74, 1.18) | 0.57 |
| PLCO ${ }^{\text {b }}$ | 1,167 | 1,093 | 0.145 | 0.137 | 1.09 (0.92, 1.30) | 0.31 |
| All combined ${ }^{\text {d }}$ | 12,102 | 16,913 | - | 0.150 | 1.21 (1.15, 1.26) | 6.2E-15 |
| $\mathbf{P}_{\text {het }}$ |  |  |  |  | 0.14 |  |
| $\mathrm{I}^{2}$ |  |  |  |  | 32.9 |  |
| C. Results for rs445114[T] on 8q24 |  |  |  |  |  |  |
| Iceland ${ }^{\text {a }}$ | 1,727 | 35,382 | 0.710 | 0.672 | 1.20 (1.11, 1.29) | 5.0E-06 |
| The Netherlands | 910 | 1,832 | 0.676 | 0.650 | 1.13 (1.00, 1.27) | 0.048 |
| Spain | 490 | 1,387 | 0.660 | 0.624 | 1.17 (1.01, 1.36) | 0.041 |
| ACS | 1,757 | 1,768 | 0.651 | 0.618 | 1.15 (1.05, 1.27) | 4.0E-03 |
| AtBC | 925 | 919 | 0.702 | 0.661 | 1.22 (1.06, 1.40) | 6.6E-03 |
| FPCC | 655 | 655 | 0.647 | 0.635 | 1.05 (0.90, 1.24) | 0.52 |
| HPFS | 595 | 608 | 0.613 | 0.633 | 0.91 (0.77, 1.07) | 0.26 |
| PLCO | 1,175 | 1,100 | 0.641 | 0.618 | 1.13 (1.00, 1.28) | 5.7E-02 |
| All combined ${ }^{\text {d }}$ | 8,234 | 43,651 | - | 0.639 | 1.14 (1.10, 1.19) | 4.7E-10 |
| $\mathbf{P}_{\text {het }}$ |  |  |  |  | 0.13 |  |
| $\mathrm{I}^{2}$ |  |  |  |  | 37.0 |  |
| D. Results for rs8102476[C] on 19q13.2 |  |  |  |  |  |  |
| Iceland ${ }^{\text {a }}$ | 1,941 | 35,330 | 0.517 | 0.495 | 1.09 (1.03, 1.17) | 6.4E-03 |
| Chicago, Illinois | 1,086 | 1,172 | 0.612 | 0.579 | 1.15 (1.02, 1.29) | 0.024 |
| Finland | 2,629 | 1,739 | 0.481 | 0.435 | 1.21 (1.11, 1.31) | $2.1 \mathrm{E}-05$ |
| The Netherlands | 1,086 | 1,830 | 0.567 | 0.528 | 1.17 (1.05, 1.30) | 4.2E-03 |
| Nashville, Tennessee | 596 | 689 | 0.565 | 0.553 | 1.05 (0.90, 1.23) | 0.55 |
| Spain | 728 | 1,389 | 0.641 | 0.619 | 1.10 (0.96, 1.25) | 0.16 |

Association results from a GWAS in Iceland for variants reported to confer risk of prostate cancer identified through GWAS

| Marker, [risk allele] and (correlated marker(s) ${ }^{a}$ | Locus | Cases$(\mathbf{N})$ | Controls <br> (N) | Frequency |  | OR (95\% CI) | P-value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Cases | Controls |  |  |
| rs2710646 [A], (rs721048) ${ }^{6}$ | 2p15 | 1,882 | 35,145 | 0.224 | 0.203 | 1.14 (1.05, 1.24) | $2.5 \times 10^{-3}$ |
| rs $2660753[\mathrm{~T}]^{4}$ | 3p12 | 1,725 | 35,362 | 0.110 | 0.100 | 1.11 (0.99, 1.25) | 0.075 |
| rs401681 [C] ${ }^{17}$ | 5p15 | 1,962 | 35,400 | 0.562 | 0.547 | 1.07 (1.00, 1.14) | 0.066 |
| rs $9364554[\mathrm{~T}]^{4}$ | 6 q 25 | 1,725 | 35,399 | 0.322 | 0.309 | 1.06 (0.99, 1.15) | 0.11 |
| rs 10486567 [G] $]^{5}$ | 7 p 15 | 1,725 | 35,392 | 0.787 | 0.765 | 1.13 (1.04, 1.24) | $4.4 \times 10^{-3}$ |
| rs6465657 [C] ${ }^{4}$ | 7 q 21 | 1,724 | 35,358 | 0.432 | 0.421 | 1.04 (0.97, 1.12) | 0.26 |
| rs $1447295[\mathrm{~A}]^{9}$ | 8 q 24 (1) | 1,821 | 35,470 | 0.165 | 0.111 | 1.58 (1.43, 1.74) | $2.2 \times 10^{-19}$ |
| $\text { rs } 16901979[A]^{1}$ | 8 q 24 (2) | 1,726 | 35,403 | 0.073 | 0.042 | 1.80 (1.55, 2.09) | $2.5 \times 10^{-14}$ |
| rs6983267 [G] ${ }^{7}$ | 8q24 (3) | 1,724 | 35,367 | 0.581 | 0.551 | 1.13 (1.05, 1.22) | $7.5 \times 10^{-4}$ |
| rs1571801 [A] ${ }^{15}$ | $9 \mathrm{q} 33{ }^{b}$ | 1,721 | 35,303 | 0.261 | 0.276 | 0.93 (0.85, 1.01) | 0.068 |
| rs10993994 [T] ${ }^{4,5}$ | 10q11 | 1,727 | 35,397 | 0.410 | 0.384 | 1.11 (1.04, 1.20) | $3.7 \times 10^{-3}$ |
| rs4962416 [C] ${ }^{5}$ | $10 \mathrm{q} 26^{\text {c }}$ | 1,724 | 35,322 | 0.223 | 0.221 | 1.02 (0.94, 1.11) | 0.68 |
| rs 10896450 [G], (rs 108964495 ${ }^{\text {a }}$, rs793 13424 ${ }^{\text {4 }}$ | 11 q 13 | 1,951 | 35,394 | 0.501 | 0.469 | 1.13 (1.06, 1.21) | $2.5 \times 10^{-4}$ |
| rs $4430796[\mathrm{~A}]^{3}$ | 17q12 | 1,726 | 35,397 | 0.559 | 0.517 | 1.19 (1.10, 1.28) | $8.3 \times 10^{-6}$ |
| rs 11649743 [G] ${ }^{18}$ | 17 q 12 | 1,747 | 35,405 | 0.812 | 0.799 | 1.09 (0.99, 1.19) | 0.066 |
| rs $1859962[\mathrm{G}]^{3}$ | 17q24.3 | 1,746 | 35,124 | 0.493 | 0.455 | 1.16 (1.08, 1.25) | $3.7 \times 10^{-5}$ |
| rs2735839 [G] ${ }^{4}$ | 19 q 13.33 | 1,726 | 35,376 | 0.879 | 0.865 | 1.14 (1.02, 1.27) | 0.021 |
| rs9623117 [C] ${ }^{19}$ | $22 \mathrm{q} 133^{b}$ | 1,724 | 35,389 | 0.208 | 0.208 | 1.00 (0.91, 1.10) | 0.99 |
| rs5945572 [A] ${ }^{6}$, (rs5945 6194) | Xp11 | 1,899 | 35,384 | 0.416 | 0.369 | 1.22 (1.11, 1.34) | $6.1 \times 10^{-5}$ |

[^1]Table 3
Population distribution in Iceland of ORs for 22 prostate cancer susceptibility variants

| OR-range | Population percentage |
| :---: | :---: |
| $<0.5$ | $9.5 \%$ |
| $0.5-0.75$ | $25.2 \%$ |
| $0.75-1$ | $24.7 \%$ |
| $1-1.5$ | $27.6 \%$ |
| $1.5-2$ | $9.1 \%$ |
| $2-2.5$ | $2.7 \%$ |
| $>2.5$ | $1.3 \%$ |

Results from a multi-variant risk model analysis for prostate cancer in Iceland based on susceptibility variants in tables 1 and 2. Results from Iceland were used for all variants in table 1 and 2, except rs1571801 on 9q33 since its effect was in the opposite direction, and rs 10896450 on 11 q 13 for which data for the refinement SNP in table 1 was used. Odds ratios (OR) were calculated for all possible genotype combinations based on 22 variants and expressed relative to the average general population risk, assuming the multiplicative model between variants. The combined OR estimates were then divided into OR-ranges and presented along with the percentage of the population within each OR-range. The general population risk was determined using a frequency-weighted average risk for all possible genotypes.


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    AUTHOR CONTRIBUTION
    The study was designed and results were interpreted by J.G., P.S., A.K., U.T., T.R., and K.S. Statistical analysis was carried out by P.S., D.F.G., J.G., M.L.F., and A.K. Subject recruitment, biological material collection and handling along with genotyping was supervised and carried out by J.G., B.A., K.R.B., D.N.M., G.O., M.J., S.N.S., A.S., T.W., T.T., J.P.B., K.M.Mc., K.M.B., B.S., J.Godino, S.N., F.F., L.M., E.P., K.K.A., I.M.vO., B.K.S., B.T.H., D.K., C.Z., K.K., J.R.G., G.V.E., E.J., W.J.C., J.I.M., L.A.K., J.R.S., J.S., R.B.B., U.T. and T.R. Authors J.G., P.S., D.F.G., T.R. and K.S. drafted the manuscript. All authors contributed to the final version of the paper. Principal investigators and corresponding authors for the respective replication study populations are: The Netherlands, L.A.K.; Spain, J.I.M.; Chicago, W.J.C.; Nashville, J.R.S.; Finland, J.S.
    COMPETING INTERESTS STATEMENT
    The authors from deCODE genetics declare competing financial interests.
    URLs
    Cancer Genetics Markers of Susceptibility (CGEMS) study (http://cgems.cancer.gov/) The Icelandic Cancer Registry (http:// www.krabbameinsskra.is/indexen.jsp?icd=C61) The International HapMap Project (http://hapmap.org/index.html) Cancer stat fact sheets at NORDCAN, The Association of the Nordic Cancer Registries (http://www-dep.iarc.fr/NORDCAN/english/ StatsFact.asp?cancer=241\&country=352)

[^1]:    ${ }^{a}$ Shown in the table are GWAS from Iceland for variants that have been identified through GWAS results (published up to February 2009) and the original publication(s). Highly correlated markers are shown in parenthesis as well as the study reporting them. All P values are two-sided. Shown are the corresponding numbers of cases and controls ( N ), allelic frequencies of variants in affected and control individuals, the allelic odds-ratio (OR) with $95 \%$ confidence interval $(95 \% \mathrm{CI})$ and $P$ value adjusted for relatedness.
    ${ }^{b}$ The original results published for the loci on 9 q 3315 and 22 q 1319 were from a study on cases with aggressive prostate cancer. Results for these two loci in Icelandic cases ( $\mathrm{N}=693$ ) with more aggressive prostate cancer (Gleason score $>6$ and/or T3 or higher and/or node positive and/or metastatic disease), using the same set of controls, were not significant (rs 1571801; OR $\mathrm{aggr}=0.90$ and $\mathrm{P}=0.080$, rs9623117; $\mathrm{OR}_{\text {aggr }}=1.00$ and $\mathrm{P}=0.94$ ).

