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A Randomized Algorithm for the Capacity of Finite-State Channels

Guangyue Han

Abstract—Inspired by ideas from the field of stochastic approximation, we propose a randomized algorithm to compute the capacity of a finite-state channel with a Markovian input. When the mutual information rate of the channel is concave with respect to the chosen parameterization, the proposed algorithm proves to be convergent to the capacity of the channel almost surely with the derived convergence rate. We also discuss the convergence behavior of the algorithm without the concavity assumption.

Index Terms—Finite-state channel, memory channel, capacity.

I. INTRODUCTION

DISCRETE-TIME finite-state channels are a broad class of channels which have attracted plenty of interest in information theory; prominent examples of such channels include partial response channels [52], [57], Gilbert-Elliott channels [21], [44] and noisy input-restricted channels [63], which are widely used in a variety of real-life applications, such as magnetic and optical recording [43], and communications over channels with inter-symbol interference [19]. The computation of the capacity of a finite-state channel is notoriously difficult and has been open for decades. For a discrete memoryless channel with a discrete memoryless source at its input, the classical Blahut-Arimoto algorithm (BAA) [3], [14] can effectively compute the channel capacity, however, for almost all nontrivial finite-state channels, little is known about the channel capacity other than some bounds (see [6], [53], [63] and references therein), which are numerically computed using Monte Carlo approaches. The methods in these work are believed to produce fairly precise numerical approximations of the capacity of certain classes of finite-state channels, however there are no general proofs to justify such beliefs.

Recently, Vontobel *et al.* have proposed a generalized Blahut-Arimoto algorithm (GBAA) [60] to maximize the mutual information rate of a finite-state machine channel with

a finite-state machine source at its input. This interesting algorithm has attracted a great deal of attention due to the observations that it fairly precisely approximates the channel capacity for a number of practical channels. (Notably, some results that were derived in the context of the GBAA have proven to be useful for analyzing the Bethe entropy function of some graphical models that appear in the context of low-density parity-check codes [58] and for approximately computing the permanent of a non-negative matrix [59].) For a finite-state channel, let X denote the input Markov process and Y its corresponding output process, which, by definition, is a *hidden Markov process* [15]. In contrast to the BAA, the proof of the convergence of the GBAA depends on the extra assumption that $I(X; Y)$ and $H(X|Y)$ are both concave with respect to a chosen parameterization, which has been posed as Conjecture 74 in [60]. Example 1, however, shows that the concavity conjecture is not true in general; for other examples showing $I(X; Y)$ and $H(X|Y)$ fail to be concave, see [37]. Here, we remark that the validity of Conjecture 74 is only sufficient to prove the convergence of the GBAA, which, however, may also converge even if the conjecture does not hold.

One of the hurdles encountered in computing the finite-state channel capacity is the problem of optimizing $I(X; Y)$ for a given parameterization, which requires in-depth understanding of the behavior of $H(Y)$ as a function of the given parameters. More specifically, there has long been a lack of understanding on the following two issues:

- (I) How to effectively compute the entropy rate of hidden Markov processes?
- (II) How does the entropy rate of hidden Markov processes vary as the underlying Markov processes and the channels vary?

As elaborated below, recently, these two issues have been partially addressed by the information theory community.

Related Work on (I): It is well known that $H(X)$ has a simple analytic formula; in stark contrast, there is no simple and explicit formula of $H(Y)$ for most non-degenerate channels ever since hidden Markov processes (or, more precisely, hidden Markov models) were formulated more than half a century ago. Here, we remark that Blackwell [12] showed that $H(Y)$ can be written as an integral of an explicit function on a simplex with respect to the Blackwell Measure. However, the Blackwell measure seems to be rather complicated for effective computation of $H(Y)$. Since 2000, there has been a rebirth of interest in computing and estimating $H(Y)$ in a variety of scenarios: the Blackwell measure has been

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used to bound $H(Y)$ [46], a variation on the classical Birch bounds [11] can be found in [18] and a new numerical approximation of $H(Y)$ has been proposed in [42]. Generalizing Blackwell's idea, an integral formula for the derivatives of $H(Y)$ has been derived in [51].

The celebrated Shannon-McMillan-Breiman theorem states that the n -th order *sample entropy* $-\log p(Y_1^n)/n$ converges to $H(Y)$ almost surely. Based on this, efficient Monte Carlo methods for approximating $H(Y)$ were proposed independently by Arnold and Loeliger [5], Pfister *et al.* [49], Sharma and Singh [54]. However, more quantitative descriptions of the convergence behavior of the proposed methods, such as rate of convergence, asymptotic normality and so on, are lacking in these work. Recently, a central limit theorem (CLT) [50] for the sample entropy has been derived as a corollary of a CLT for the top Lyapunov exponent of a product of random matrices; a functional CLT has also been established in [33]. To some extent, these two CLTs suggested that the Monte Carlo methods are "accurate" for the purpose of approximating $H(Y)$ in the sense that, roughly speaking, the methods are more likely to yield values concentrated around $H(Y)$. There are also other related work in different contexts, such as [31], [32], and [35].

Recently, we have obtained [23] a number of limit theorems for the sample entropy of Y . These limit theorems can be viewed as further refinements of the Shannon-McMillan-Breiman theorem, which is the backbone of information theory. More specifically, [23, Th. 1.2] is a CLT with an error-estimate, which can be used to characterize the rate of convergence of the Monte Carlo methods in [5], [49], and [54], and [23, Th. 1.5] is a large deviation result, which gives a sub-exponential decaying upper bound on the probability of the sample entropy $-\log p(Y_1^n)/n$ deviating from $H(Y)$. Among many other applications, such as deriving non-asymptotic coding theorems [62], these theorems positively confirmed the effectiveness of using the Shannon-McMillan-Breiman theorem to approximate $H(Y)$.

Related Work on (II): The behavior of $H(Y)$ (as a function of the underlying Markov chain and the channel) is of significance in a number of scientific disciplines; particularly in information theory, it is of great importance for computing/estimating the capacity of finite-state channels. However, some of the basic problems, such as smoothness (or even differentiability) of $H(Y)$, have long remained unknown. Recently, asymptotical behavior of $H(Y)$ has been studied in [4], [34], [45]–[48], [51], [64], and [65]. Particularly in [64], for a special type of hidden Markov chain Y , the Taylor series expansion of $H(Y)$ is given under the assumption that $H(Y)$ is analytic. Under mild assumptions, analyticity of $H(Y)$ has been established in [25]; see also related work in [2], [15], [42], [51], [64], [65], and references therein. The framework in [25] has been generalized to continuous-state settings and further provides useful tools and techniques for our subsequent work, such as derivatives [26], asymptotics [27], and concavity [28] of $H(Y)$.

Equipped with ideas and techniques from the above-mentioned work on (I) and (II), we are more prepared to make further progress towards the computation of

the channel capacity. In particular, the ideas and techniques in [23] and [25] are vital to this paper. Roughly speaking, [25] proves that the entropy rate of hidden Markov chains is a "nicely behaved" function; and [23] confirms that it can be "well-approximated" using Monte Carlo simulations. The simulator of the derivative of $I(X; Y)$ as specified in Section IV, which is crucial to this work, is an "offspring" of the two schools of thoughts in [23] and [25].

Stochastic approximation methods refer to a family of recursive stochastic algorithms, aiming to find zeroes or extrema of functions whose values can only be estimated via noisy observations. The extensive literature on stochastic approximation has grown up around two prototypical algorithms, the Robbins-Monro algorithm and the Kiefer-Wolfowitz algorithm, mainly concerning the convergence analysis on these two algorithms and their variants; we refer the reader to [36] for an exposition to the vast literature on stochastic approximation.

Inspired by ideas in stochastic approximation, we propose a randomized algorithm to compute the capacity of a class of finite-state channels with input Markov processes supported on some mixing finite-type mixing constraint. Bearing the same spirit as the Robbins-Monro algorithm and the Kiefer-Wolfowitz algorithm, the proposed algorithm, in many subtle respects, differs from both of them. The main task of this paper is to conduct a convergence analysis of the proposed algorithm, which employs some established ideas and techniques from the field of stochastic approximation. In particular, the proofs in Section VIII are largely inspired by [56], which has credited origins of some of its techniques to earlier work, such as [8], [36], and [39]. However, neither the results nor the proofs in [56] or any of previous work imply our results; as a matter of fact, considerable amount of simplification and adaptation of the techniques in [56] have been incorporated into this work.

Although described in different languages, our settings are essentially the same as in [60]. On the other hand, as opposed to the GBAA, the concavity of $I(X; Y)$ alone is already sufficient to guarantee the convergence of our algorithm. Here, let us note that that for certain classes of channels (see Example 1), $I(X; Y)$ is indeed concave with respect to certain parameterization, whereas $H(X|Y)$ fails to be concave with respect to the same parameterization.

Characterizing the maximal rate at which the information can be transmitted through a given channel, the capacity is the most fundamental notion in information theory. The capacity-achieving (input) distribution will further provide us insightful guidance towards designing coding schemes that actually achieve the promised capacity. Apparently, our proposed algorithm would be of fundamental significance to both information theoretic research and practical applications to tele-communications and data storage.

The organization of the paper is as follows. We first describe our channel model in greater detail in Section II and we then present our algorithm in Section III. In Section IV, we propose a simulator for the derivative of $I(X; Y)$ and discuss its convergence behavior. The convergence of the algorithm is established in Section V, while the rate of convergence of the

algorithm with and without concavity conditions are derived in Sections VII and VIII, respectively. In Section IX, we discuss the capacity-achieving distribution of a special class of finite-state channels. Our proposed algorithm has been implemented for a variety of input-constrained finite-state channels, and the results of extensive numerical simulations for these channels will be presented in Section X.

II. CHANNEL MODEL

In this section, we specify the channel model considered in this paper in greater detail, which is essentially the same as the one considered in [60].

Let \mathcal{X} be a finite alphabet and let

$$\mathcal{X}^2 = \{(i, j) : i, j \in \mathcal{X}\}.$$

Let Π denote the set of all stationary irreducible first-order Markov chains over the alphabet \mathcal{X} . For a given subset $F \subset \mathcal{X}^2$, define

$$\Pi_F = \{X \in \Pi : X_{i,j} = 0, (i, j) \in F\},$$

where we have identified an irreducible first-order Markov chain with its transition probability matrix (so, $X_{i,j}$ represents the probability of X transiting from state i to state j). Furthermore, for any $\epsilon > 0$, define

$$\Pi_{F,\epsilon} = \{X \in \Pi_F : X_{i,j} \geq \epsilon, (i, j) \notin F\}.$$

Obviously, if some $X \in \Pi_{F,\epsilon}$ is primitive (namely, irreducible and aperiodic), then any $X' \in \Pi_{F,\epsilon}$ is primitive; in this case, we say F is a *mixing* finite-type constraint. Here, let us note that a mixing finite-type constraint can be defined in a much more general context [40].

The motivation for considering finite-type constraints mainly comes from magnetic recording, where input sequences are required to satisfy certain mixing finite-type constraints in order to eliminate the most damaging error events [43]. The best known example is the so-called (d, k) -RLL constraint over the alphabet $\{0, 1\}$, which forbids any sequence with fewer than d or more than k consecutive zeros in between two successive 1's.

In this paper, we are concerned with a discrete-time finite-state channel with some input constraint. Let X, Y, S denote the channel input, output and state processes over finite alphabets \mathcal{X}, \mathcal{Y} and \mathcal{S} , respectively. Assume that

- (II.a) For some mixing finite-type constraint $F \subset \mathcal{X}^2$ and some $\epsilon > 0$, $X \in \Pi_{F,\epsilon}$.
- (II.b) (X, S) is a first-order stationary Markov chain whose transition probabilities satisfy

$$p(x_n, s_n | x_{n-1}, s_{n-1}) = p(x_n | x_{n-1})p(s_n | x_n, s_{n-1}),$$

where $p(s_n | x_n, s_{n-1}) > 0$ for any s_{n-1}, s_n, x_n .

- (II.c) the channel is stationary, which is characterized by

$$p(y_n | y_1^{n-1}, x_1^n, s_1^{n-1}) = p(y_n | x_n, s_{n-1}) > 0,$$

that is, conditioned on the pair x_n, s_{n-1} , the output y_n is statistically independent of all outputs, inputs and states prior to y_n, x_n and s_{n-1} , respectively.

It will be clear later (see Remark 5) that the channel specified as above is indecomposable [20], and its capacity can be computed as

$$C_F = \sup I(X; Y) = \sup \lim_{n \rightarrow \infty} I_n(X; Y),$$

where the supremum is over all X satisfying Assumptions (II.a) and

$$I_n(X; Y) \triangleq \frac{H(X_1^n) + H(Y_1^n) - H(X_1^n, Y_1^n)}{n}. \quad (1)$$

The fact that Y and (X, Y) are both hidden Markov processes makes it apparent that solutions to (I) and (II) are essential for computing C_F .

Assume that $\Pi_{F,\epsilon}$ is analytically parameterized by $\theta \in \Theta \subset \mathbb{R}^d$, $d \geq 1$, where Θ denote the entire parameter space. Then, naturally, $X = X(\theta)$ and $Y = Y(\theta)$ are also analytically parameterized by θ . Under this parameterization, we would like to find $\theta^* \in \Theta$ that maximizes $I(X(\theta); Y(\theta))$.

Remark 1: The assumption that X is a first-order Markov chain is for notional convenience only: through a usual “reblocking” technique, the case for a Markov chain of any order can be boiled down to the first-order case.

Remark 2: One natural goal is to find $X \in \Pi_F$ that maximizes $I(X; Y)$. However, in this paper, we will restrict our attention to $\Pi_{F,\epsilon}$ for a given $\epsilon > 0$; such restriction will be justified in Section IX.

III. THE ALGORITHM

For a given $1/2 < a < 1$, choose the so-called step sizes

$$a_n = \frac{1}{n^a}, \quad n = 1, 2, \dots; \quad (2)$$

apparently, $\{a_n\}$ satisfies

$$\sum_{n=0}^{\infty} a_n = \infty, \quad \sum_{n=0}^{\infty} a_n^2 < \infty,$$

which are the typical conditions imposed on step sizes in a generic stochastic approximation method. Letting A_n denote the event “ $\theta_n + a_n g_n^b(\theta_n) \notin \Theta$ ”, we propose to find θ^* through the following recursive procedure:

$$\theta_{n+1} = \begin{cases} \theta_n, & \text{if } A_n \text{ occurs,} \\ \theta_n + a_n g_n^b(\theta_n), & \text{otherwise;} \end{cases} \quad (3)$$

here $b > 0$, the initial θ_0 is randomly selected from Θ , and $g_n^b(\theta)$ is a to-be-specified simulator (see Section IV) for $I'(X(\theta); Y(\theta))$, where $'$ means the derivative taken with respect to θ . Throughout the paper, we assume that

$$0 < \beta < \alpha < 1/3, \quad 2a + b - 3b\beta > 1; \quad (4)$$

here, α, β are some “hidden” parameters involved in the definition of $g_n^b(\theta)$, which will be defined in Section IV.

Remark 3: The idea of the algorithm is very simple and can be roughly explained as follows: at time n , the derivative $I'(X(\theta_n); Y(\theta_n))$, which can be “well-approximated” by $g_n^b(\theta_n)$, indicates whether the function $I(X(\theta), Y(\theta))$ is monotone increasing or not at θ_n .

Accordingly, while ensuring all θ_n stay within the region Θ , θ_n will be updated to θ_{n+1} , most likely giving rise to a larger value of $I(X(\theta_{n+1}), Y(\theta_{n+1}))$. The subtlety is that one has to appropriately choose step sizes $\{a_n\}$: too small step sizes will cause a premature convergence of $\{\theta_n\}$, while too large step sizes will likely lead to an eventual divergence of $\{\theta_n\}$.

Remark 4: In our notation, one of the key differences between our algorithm and the GBAA in [60] is that our algorithm is meant to find θ^* that maximizes the mutual information rate $I(X(\theta); Y(\theta))$, whereas the GBAA to find θ_n^* that maximizes the mutual information $I_n(X(\theta); Y(\theta))$ for some arbitrary yet fixed n .

IV. A SIMULATOR OF $I'(X; Y)$

As stated in Section I, albeit rather difficult to compute analytically, $I_n(X; Y)$ can be well-approximated via Monte Carlo simulations. In this section, we propose a simulator for $I'(X; Y)$. Needless to say, an effective simulator guaranteeing an “accurate” approximation to $I'(X; Y)$ is crucial to our algorithm. To some extent, our simulator is inspired by Bernstein’s blocking method [9], which is a well-established tool in proving limit theorems for mixing sequences [16].

Now, consider a stationary stochastic process $Z = Z_{-\infty}^{\infty}$ satisfying the following assumptions:

(IV.a) There exist strictly positive constants $\hat{C}_1 < \hat{C}_2$ such that for all z_{-n}^0 ,

$$\hat{C}_1 \leq p(z_0|z_{-n}^{-1}) \leq \hat{C}_2.$$

(IV.b) There exist $C > 0$, $0 < \lambda < 1$ such that for all n ,

$$\begin{aligned} \psi_Z(n) \triangleq & \sup_{U \in \mathcal{B}(Z_{-\infty}^{\infty}), V \in \mathcal{B}(Z_0^{\infty}), P(U), P(V) > 0} \\ & \times \frac{|P(V|U) - P(V)|}{P(V)} \leq C\lambda^n, \end{aligned}$$

where $\mathcal{B}(Z_i^j)$ denotes the σ -field generated by $\{Z_k : k = i, i+1, \dots, j\}$.

(IV.c) There exist $C > 0$, $0 < \rho < 1$ such that for any two $z_{-m}^0, \hat{z}_{-\hat{m}}^0$ with $z_{-n}^0 = \hat{z}_{-\hat{m}}^0$ (here $m, \hat{m} \geq n \geq 0$),

$$|p(z_0|z_{-m}^{-1}) - p(\hat{z}_0|\hat{z}_{-\hat{m}}^{-1})| \leq C\rho^n.$$

Remark 5: Conditions (IV.a)-(IV.c) are the same ones used in [23, Sec. 2], which are essential for establishing the main results in [23]. As observed in [23], Conditions (II.a)-(II.c) imply that Y and (X, Y) both satisfy Conditions (IV.a)-(IV.c), and moreover, there exist $C > 0$, $0 < \rho < 1$ such that for any two initial states s'_0, s''_0 ,

$$|p(s_n|x_1^n, s'_0) - p(s_n|x_1^n, s''_0)| \leq C\rho^n.$$

This implies that, as claimed in Section II, the finite-state channels considered in this paper are indecomposable.

Now, for $0 < \beta < \alpha < 1/3$, define

$$q = q(n) \triangleq n^\beta, \quad r = r(n) \triangleq n^\alpha, \quad k = k(n) \triangleq n/(n^\alpha + n^\beta),$$

where, to avoid notational cumbersomeness, we have omitted rounding operators acting on n^β , n^α and $n/(n^\alpha + n^\beta)$,

which are not necessarily integers. For any j with $iq + (i-1)r + 1 \leq j \leq ir + ip$, define

$$W_j = W_j(Z_{j-\lfloor q/2 \rfloor}^j) \triangleq -\frac{p'(Z_{j-\lfloor q/2 \rfloor}^j)}{p(Z_{j-\lfloor q/2 \rfloor}^j)} \log p(Z_j|Z_{j-\lfloor q/2 \rfloor}^{j-1}), \quad (5)$$

and furthermore

$$\zeta_i \triangleq W_{iq+(i-1)r+1} + \dots + W_{iq+ir}, \quad S_n \triangleq \sum_{i=1}^{k(n)} \zeta_i. \quad (6)$$

Remark 6: The rough idea behind the definition of S_n is as follows: Note that each ζ_i only involves all Z_j with $j \in I_i \triangleq \{iq+(i-1)r+1-\lfloor q/2 \rfloor, \dots, iq+ir\}$, and for any two $j_1 \in I_{i_1}$, $j_2 \in I_{i_2}$, where i_1, i_2 are distinct, we have $|j_1 - j_2| \geq \lfloor q/2 \rfloor$. It then follows from Condition (IV.b) that for large enough n , all ζ_i are “weakly dependent”, and furthermore, the arithmetic mean of all ζ_i will converge to the mean of ζ_i , which can be verified to be close to $H'(Z)$.

Now, we are ready to define our simulator for $I'(X; Y)$.

Definition 1:

$$\begin{aligned} g_n &= g_n(X_1^n, Y_1^n) \\ &\triangleq H'(X_2|X_1) + S_n(Y_1^n)/(kr) - S_n(X_1^n, Y_1^n)/(kr). \end{aligned}$$

The following lemma, whose proof is somewhat similar to that of [23, Lemma 3.3], gives an estimate of the variance of S_n .

Lemma 1: For Z satisfying Conditions (IV.a), (IV.b) and (IV.c),

$$E[(S_n - E[S_n])^2] = O(kr q^3).$$

Proof: As in the proofs of [23, Lemmas 3.1 and 3.9], using Condition (IV.a), (IV.b), we can deduce that for some $0 < \lambda < 1$,

$$\begin{aligned} E[(S_n - E[S_n])^2] &= E\left[\left(\sum_{i=1}^k \zeta_i - \sum_{i=1}^k E[\zeta_i]\right)^2\right] \\ &= kE[(\zeta_1 - E[\zeta_1])^2] + O(k^2 \lambda^{q/2}). \end{aligned}$$

So, to prove the lemma, it suffices to prove that

$$E[(\zeta_1 - E[\zeta_1])^2] = O(rq^3).$$

Note that

$$E[(\zeta_1 - E[\zeta_1])^2] = \sum_{j_1, j_2=1}^r E[(W_{j_1} - E[W_{j_1}])(W_{j_2} - E[W_{j_2}])]. \quad (7)$$

It follows from Condition (IV.a) that for any j , $W_j - E[W_j] = O(q)$, and therefore, for any j_1, j_2 , we have

$$E[(W_{j_1} - E[W_{j_1}])(W_{j_2} - E[W_{j_2}])] = O(q^2). \quad (8)$$

Furthermore, similarly as in the proof of Part 2 of [23, Lemma 3.1], one verifies, using Condition (IV.a), (IV.b), that when $|j_1 - j_2| > \lfloor q/2 \rfloor$,

$$E[(W_{j_1} - E[W_{j_1}])(W_{j_2} - E[W_{j_2}])] = O(q^2 \lambda^{|j_1 - j_2| - \lfloor q/2 \rfloor}). \quad (9)$$

Combining (7), (8) and (9), we then have

$$\begin{aligned} & E[(\zeta_1 - E[\zeta_1])^2] \\ &= \left(\sum_{|j_1 - j_2| \leq \lfloor q/2 \rfloor} + \sum_{|j_1 - j_2| > \lfloor q/2 \rfloor} \right) E[(W_{j_1} - E[W_{j_1}]) \\ & \quad \times (W_{j_2} - E[W_{j_2}])] \\ &= O(rq^3). \end{aligned}$$

The proof is then complete. \square

The following three theorems characterize the performance of our simulator from different perspectives.

Using similar techniques as in [25, Proof of Theorem 1.1], the first theorem shows that on average, our simulator sub-exponentially converges to $I'(X; Y)$.

Theorem 1: For some $0 < \rho_0 < 1$, we have

$$E[g_n(X_1^n, Y_1^n)] - I'(X; Y) = O(\rho_0^{\lfloor q/2 \rfloor}).$$

Proof: Notice that for the Markov chain X , we have

$$H(X) = H(X_2|X_1).$$

So, by Remark 5, it suffices to prove that for any Z satisfying Conditions (IV.a)-(IV.c), we have

$$\frac{E[S_n]}{kr} - H'(Z) = O(\rho_1^{\lfloor q/2 \rfloor}),$$

for some $0 < \rho_1 < 1$.

Note that for any j with $iq + (i-1)r + 1 \leq j \leq iq + ir$, we have

$$\begin{aligned} E[W_j] &= - \sum_{z_{j-\lfloor q/2 \rfloor}^j} p(z_{j-\lfloor q/2 \rfloor}^j) \frac{p'(z_{j-\lfloor q/2 \rfloor}^j)}{p(z_{j-\lfloor q/2 \rfloor}^j)} \log p(z_j | z_{j-\lfloor q/2 \rfloor}^{j-1}) \\ &= - \sum_{z_{j-\lfloor q/2 \rfloor}^j} p'(z_{j-\lfloor q/2 \rfloor}^j) \log p(z_j | z_{j-\lfloor q/2 \rfloor}^{j-1}). \end{aligned}$$

Then, as in [25, Proof of Theorem 1.1], we can prove that for any small $\varepsilon > 0$,

$$\sum_{z_1^n} |p'(z_n | z_1^{n-1})| = O((1 + \varepsilon)^n).$$

This, together with Condition (IV.c), implies that for some $0 < \rho_1 < 1$,

$$E[W_j] - H'(Z) = O(\rho_1^{\lfloor q/2 \rfloor}),$$

which further implies that for some $0 < \rho_1 < 1$

$$\begin{aligned} \frac{E[S_n]}{kr} - H'(Z) &= \frac{E[S_n] - krH'(Z)}{kr} \\ &= \frac{\sum_j (E[W_j] - H'(Z))}{kr} \\ &= O(\rho_1^{\lfloor q/2 \rfloor}). \end{aligned}$$

\square

The following large deviation type lemma gives a sub-exponentially decaying upper bound on the tail probability of $g_n(X_1^n, Y_1^n)$ deviating from $I'(X; Y)$.

Theorem 2: For any $\varepsilon > 0$, there exist some $0 < \gamma, \delta < 1$ such that

$$P(|g_n(X_1^n, Y_1^n) - I'(X; Y)| \geq \varepsilon) \leq \gamma n^\delta.$$

Proof: By Lemma 1 and Remark 5, it suffices to prove that for any Z satisfying Conditions (IV.a)-(IV.c) and for any $\varepsilon > 0$, there exist $0 < \gamma, \delta < 1$ such that

$$P\left(\left|\frac{S_n - E[S_n]}{kr}\right| \geq \varepsilon\right) \leq \gamma n^\delta. \quad (10)$$

By the Markov inequality, we have

$$\begin{aligned} P(S_n - E[S_n] \geq kr\varepsilon) &= P\left(\frac{t(S_n - E[S_n])}{r^2} \geq \frac{tkr\varepsilon}{r^2}\right) \\ &\leq \frac{E[e^{t(S_n - E[S_n])/r^2}]}{e^{tk\varepsilon/r}}. \end{aligned} \quad (11)$$

As in [23, Proof of Lemma 3.9], applying Conditions (IV.a) and (IV.b), we then have

$$\begin{aligned} & E[e^{t(S_n - E[S_n])/r^2}] \\ &= E[e^{t \sum_{i=1}^{k-1} (\zeta_i - E[\zeta_i])/r^2} e^{t(\zeta_k - E[\zeta_k])/r^2}] \\ &= (1 + O(\lambda^{q(n)/2})) E[e^{t \sum_{i=1}^{k-1} (\zeta_i - E[\zeta_i])/r^2}] E[e^{t(\zeta_k - E[\zeta_k])/r^2}], \end{aligned} \quad (12)$$

for some $0 < \lambda < 1$, where we have used the fact, as explained in Remark 6, that for any two $j_1 \in I_{i_1}$, $j_2 \in I_{i_2}$, where i_1, i_2 are distinct, we have $|j_1 - j_2| \geq \lfloor q/2 \rfloor$. A recursive application of (12) yields that for any $0 < t < 1$

$$\begin{aligned} E[e^{t(S_n - E[S_n])/r^2}] &= E[e^{t \sum_{i=1}^k (\zeta_i - E[\zeta_i])/r^2}] \\ &= (1 + O(\lambda^{q(n)/2}))^{k-1} (E[e^{t(\zeta_1 - E[\zeta_1])/r^2}])^k, \end{aligned} \quad (13)$$

as n goes to infinity. By Condition (IV.a), we have

$$\zeta_1 - E[\zeta_1] = O(rq),$$

and thus,

$$O((\zeta_1 - E[\zeta_1])^2/r^4) = O(q^2/r^2) = o(1).$$

It then follows that for any $0 < t < 1$,

$$E[e^{t(\zeta_1 - E[\zeta_1])/r^2}] = 1 + o(1)t^2.$$

Choosing $t = n^{-(1-\alpha)/2}$, then, by (11) and (13), we deduce that

$$\begin{aligned} & P\left(\frac{S_n - E[S_n]}{kr} \geq \varepsilon\right) \\ &\leq \frac{E[e^{t(S_n - E[S_n])/r^2}]}{e^{tk\varepsilon/r}} \\ &\leq (1 + O(\lambda^{q(n)/2}))^k \frac{(1 + o(1)t^2)^{n^{1-\alpha}}}{(1 + t\varepsilon + O(1)t^2)^{n^{1-2\alpha}}} \\ &= O(e^{-n^{1/2-3\alpha/2}}). \end{aligned}$$

Noticing that $0 < \alpha < 1/3$ (and thus $1/2 - 3\alpha/2 > 0$), we conclude that for any $\varepsilon > 0$, there exists $0 < \gamma, \delta < 1$ such that

$$P\left(\frac{S_n - E[S_n]}{kr} \geq \varepsilon\right) \leq \gamma n^\delta.$$

With a parallel argument, one verifies that for any $\varepsilon > 0$, there exists $0 < \gamma, \delta < 1$ such that

$$P\left(\frac{S_n - E[S_n]}{kr} \leq -\varepsilon\right) \leq \gamma n^\delta,$$

which immediately implies (10). The proof is then complete. \square

The following theorem states that our simulator is asymptotically unbiased.

Theorem 3: With probability 1,

$$g_n(X_1^n, Y_1^n) \rightarrow I'(X; Y),$$

as n tends to ∞ .

Proof: It immediately follows from Theorem 2 and the Borel-Cantelli lemma. \square

Remark 7: Other than confirming the effectiveness of the proposed simulator g_n in Definition 1, Theorems 1, 2 and 3 are of interest in their own right: For any stationary process Z satisfying Conditions (IV.a), (IV.b) and (IV.c), the proofs of these theorems reveals that the sample path of Z can be used to effectively simulate the derivative of $H(Z)$. This can be viewed as an extension (for a special class of processes) of the Shannon-Millan-Breiman theorem, which states that the sample path of a stationary and ergodic process can be used to simulate its entropy rate. For the proposed algorithm in Section III, an alternative simulator can be defined as follows: g_n can be defined in the same way except that $\zeta_2, \zeta_3, \dots, \zeta_n$ are taken as independent copies of ζ_1 . It can be easily checked that Theorems 1, 2 and 3 also holds for the redefined g_n . As a consequence, the convergence results later in this paper also hold for the redefined g_n .

Remark 8: In our notation, the following expression has been proposed in [60] as a simulator of $I'(X; Y)$:

$$\begin{aligned} H'(X_2|X_1) & - \frac{p'(Y_1^n)}{p(Y_1^n)}(\log p(Y_1^n))/n \\ & + \frac{p'(X_1^n, Y_1^n)}{p(X_1^n, Y_1^n)}(\log p(X_1^n, Y_1^n))/n. \end{aligned}$$

Extensive numerical experiments conducted in [60] suggest that this simulator converges to $I'(X; Y)$ almost surely as n tends to infinity, however, there is no rigorous proof for the convergence.

Remark 9: For the purpose of effective computation to avoid accumulating round-off errors caused by multiplication, W_j can be rewritten as

$$\begin{aligned} W_j & = - \left(\frac{p'(Z_{j-\lfloor q/2 \rfloor})}{p(Z_{j-\lfloor q/2 \rfloor})} + \frac{p'(Z_{j-\lfloor q/2 \rfloor+1}|Z_{j-\lfloor q/2 \rfloor})}{p(Z_{j-\lfloor q/2 \rfloor+1}|Z_{j-\lfloor q/2 \rfloor})} \right. \\ & \quad \left. + \dots + \frac{p'(Z_j|Z_{j-\lfloor q/2 \rfloor}^{j-1})}{p(Z_j|Z_{j-\lfloor q/2 \rfloor}^{j-1})} \right) \log p(Z_j|Z_{j-\lfloor q/2 \rfloor}^{j-1}). \end{aligned}$$

Here, we remark that it is well known that all $p(Z_i|Z_{i-\lfloor q/2 \rfloor}^{i-1})$ (see [25, eqs. (4.6) and (4.14)]) and $p'(Z_i|Z_{i-\lfloor q/2 \rfloor}^{i-1})$ (see, e.g., the displayed equation after [26, eq. (2.5)]) can be effectively computed in a recursive manner.

Remark 10: It is well known that by the Shannon-McMillan-Breiman theorem, the following simulator of $I(X; Y)$,

$$H(X_2|X_1) - \log p(Y_1^n)/n + \log p(X_1^n, Y_1^n)/n, \quad (14)$$

converges to $I(X; Y)$ almost surely as n tends to infinity. Bearing the same spirit as the simulator of $I'(X; Y)$ in Definition 1, an alternative simulator of $I(X; Y)$ can be defined as follows: For any j with $iq + (i-1)r + 1 \leq j \leq ir + ip$, define

$$\hat{W}_j = \hat{W}_j(Z_{j-\lfloor q/2 \rfloor}^j) \triangleq -\log p(Z_j|Z_{j-\lfloor q/2 \rfloor}^{j-1}),$$

and furthermore

$$\hat{\zeta}_i \triangleq W_{iq+(i-1)r+1} + \dots + W_{iq+ir}, \quad \hat{S}_n \triangleq \sum_{i=1}^{k(n)} \hat{\zeta}_i.$$

Then, the following expression gives an alternative simulator of $I(X; Y)$:

$$\begin{aligned} \hat{g}_n & = \hat{g}_n(X_1^n, Y_1^n) \\ & \triangleq H(X_2|X_1) + \hat{S}_n(Y_1^n)/(kr) - \hat{S}_n(X_1^n, Y_1^n)/(kr). \end{aligned}$$

It can be verified that analogs of Theorems 1, 2 and 3 hold for the above alternative simulator; more precisely, following from parallel arguments, these three theorems still hold with g_n and $I'(X; Y)$ replaced with \hat{g}_n and $I(X; Y)$, respectively. Note that the sub-exponential upper bound as in Theorem 2, which in general does not hold for the simulator in (14), suggests that \hat{g}_n is more effective. Indeed, the alternative simulator \hat{g}_n has been implemented in our numerical simulations (See Section X) for the computation of $I(X; Y)$ and faster convergence of this simulator has been observed.

V. CONVERGENCE

Consider the following condition:

(V.a) $P(\cap_{k=1}^{\infty} \cup_{n=k}^{\infty} A_n) = 0$, that is, $A_n, n \in \mathbb{N}$, only occurs finitely many times,

which will be assumed throughout the convergence analysis in the paper. Particularly, in this section, assuming (V.a), we will show that $\{I(X(\theta_n); Y(\theta_n))\}$ converges almost surely. Obviously, if $\Theta = \mathbb{R}^d$, then Assumption (V.a) will be trivially satisfied and the iteration in (3) can be simply written as

$$\theta_{n+1} = \theta_n + a_n g_n^b(\theta_n). \quad (15)$$

Note that Assumption (V.a) may not hold true when Θ is bounded, though some other conditions, such as Condition (VII.a), may imply (5.a). Unless specified otherwise, we will simply assume that $\Theta = \mathbb{R}$ in all the proofs in this paper to avoid obscuring the main idea. The proofs of the same results under Assumption (V.a) follow from parallel arguments only with an increasing level of notational complexity.

Henceforth, we will write (for the definition of I_n , see (1))

$$f(\theta) = I(X(\theta); Y(\theta)), \quad f_n(\theta) = I_n(X(\theta); Y(\theta)).$$

Note that under Assumption (II.a), [25, Th. 1.1] implies that $f(\theta)$ is analytic and each of its derivatives is uniformly bounded over all $\theta \in \Theta$,

a key fact that will be implicitly used throughout the paper. Now, rewrite (15) as

$$\theta_{n+1} = \theta_n + a_n f'(\theta_n) + a_n R_n(\theta_n), \quad (16)$$

where

$$R_n(\theta_n) \triangleq g_{n^b}(\theta_n) - f'(\theta_n).$$

It can be easily verified that

$$\begin{aligned} & f(\theta_{n+1}) - f(\theta_n) \\ &= \int_0^1 f'(\theta_n + t(\theta_{n+1} - \theta_n))(\theta_{n+1} - \theta_n) dt \\ &= \int_0^1 f'(\theta_n)(\theta_{n+1} - \theta_n) dt \\ &\quad + \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt \\ &= a_n f'(\theta_n)(f'(\theta_n) + R_n(\theta_n)) \\ &\quad + \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt \\ &= a_n f'^2(\theta_n) + \hat{R}_n(\theta_n), \end{aligned} \quad (17)$$

where

$$\begin{aligned} \hat{R}_n(\theta_n) &\triangleq a_n f'(\theta_n) R_n(\theta_n) \\ &\quad + \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt. \end{aligned}$$

Lemma 2: $\sum_{n=0}^{\infty} \hat{R}_n(\theta_n)$ converges almost surely.

Proof: Let

$$T_1 = \sum_{n=0}^{\infty} a_n f'(\theta_n) R_n(\theta_n),$$

and

$$T_2 = \sum_{n=0}^{\infty} \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt.$$

It suffices to prove that T_1, T_2 both converge almost surely.

For T_1 , note that

$$\begin{aligned} T_1 &= \sum_{n=0}^{\infty} a_n f'(\theta_n)(g_{n^b}(\theta_n) - f'(\theta_n)) \\ &= \sum_{n=0}^{\infty} a_n f'(\theta_n)(g_{n^b}(\theta_n) - f'_{n^b}(\theta_n)) \\ &\quad + \sum_{n=0}^{\infty} a_n f'(\theta_n)(f'_{n^b}(\theta_n) - f'(\theta_n)). \end{aligned}$$

It follows from the proof of Theorem 1.1 (more precisely, the last displayed equation in the proof of Theorem 1.1, coupled with Cauchy's integral formula) that there exists $0 < \rho_0 < 1$ such that for all $\theta \in \Theta$,

$$|(f'_{n^b}(\theta) - f'(\theta))| \leq \rho_0^{n^b},$$

which immediately implies that

$$\sum_{n=0}^{\infty} a_n |f'(\theta_n)| |(f'_{n^b}(\theta_n) - f'(\theta_n))| \leq \sum_{n=0}^{\infty} a_n |f'(\theta_n)| \rho_0^{n^b} < \infty. \quad (18)$$

Then, using (2) and Lemma 1, one verifies that uniformly over all $\theta_n \in \Theta$,

$$\sum_{n=0}^{\infty} E[\{a_n^2 (f'(\theta_n))^2 R_n^2(\theta_n)\}] = \sum_{n=0}^{\infty} O\left(\frac{1}{n^{2a+b(1-3\beta)}}\right), \quad (19)$$

which converges since $2a + b - 3b\beta > 1$. Noting that $\{a_n f'(\theta_n) R_n(\theta_n), \mathcal{B}(X_1^n)\}$ is a Martingale difference sequence and applying Doob's Martingale convergence theorem (see [55, Th. 2.8.7]), we deduce that

$$\sum_{n=0}^{\infty} a_n f'(\theta_n)(g_{n^b}(\theta_n) - f'_{n^b}(\theta_n))$$

converges with probability 1. The almost sure convergence of T_1 then follows.

For T_2 , it is easy to check that

$$\begin{aligned} & \left| \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt \right| \\ &= O((\theta_{n+1} - \theta_n)^2) \\ &= O(a_n^2 (f'(\theta_n))^2) + O(a_n^2 R_n^2(\theta_n)), \end{aligned}$$

where the constants in the above O -terms are uniform over all θ_n, θ_{n+1} . Similarly as in deriving (18) and (19), we have

$$\sum_{n=0}^{\infty} a_n^2 (f'_{n^b}(\theta_n) - f'(\theta_n))^2 < \infty,$$

and

$$\sum_{n=0}^{\infty} E[a_n^2 (g_{n^b}(\theta_n) - f'_{n^b}(\theta_n))^2] < \infty,$$

and furthermore,

$$\sum_{n=0}^{\infty} a_n^2 (g_{n^b}(\theta_n) - f'_{n^b}(\theta_n))^2$$

converges almost surely. This, together with (18), further implies that

$$\sum_{n=0}^{\infty} a_n^2 |(g_{n^b}(\theta_n) - f'_{n^b}(\theta_n))(f'_{n^b}(\theta_n) - f'(\theta_n))|$$

converges almost surely. Recalling that

$$R_n(\theta_n) = g_{n^b}(\theta_n) - f'_{n^b}(\theta_n) + f'_{n^b}(\theta_n) - f'(\theta_n),$$

we conclude that

$$\sum_{n=0}^{\infty} a_n^2 R_n^2(\theta_n) < \infty,$$

which further implies that

$$\sum_{n=0}^{\infty} \int_0^1 (f'(\theta_n + t(\theta_{n+1} - \theta_n)) - f'(\theta_n))(\theta_{n+1} - \theta_n) dt$$

converges almost surely. The proof is then complete. \square

We are now ready for the following convergence theorem, whose proof closely follows that of [56, Lemma 7],

whose general ideas are reminiscent of the standard proof of the Martingale convergence theorem [55].

Theorem 4: With probability 1, we have

$$\lim_{n \rightarrow \infty} f'(\theta_n) = 0 \quad \text{and} \quad \lim_{n \rightarrow \infty} f(\theta_n) \text{ exists.}$$

Proof: Recall that

$$f(\theta_{n+1}) - f(\theta_n) = a_n f'^2(\theta_n) + \hat{R}_n(\theta_n),$$

a recursive application of which implies

$$f(\theta_n) = f(\theta_0) + \sum_{i=0}^{n-1} a_i (f'(\theta_i))^2 + \sum_{i=0}^{n-1} \hat{R}_i(\theta_i).$$

Applying Lemma 2, we deduce that with probability 1,

$$\sum_{i=0}^{\infty} a_i (f'(\theta_i))^2 < \infty,$$

which, in turn, implies that $\lim_{n \rightarrow \infty} f(\theta_n)$ exists and furthermore there is a subsequence $\{\theta_{n_j}\}$ such that $f'(\theta_{n_j})$ converges to 0 as j tends to infinity.

We now prove that

$$\lim_{n \rightarrow \infty} f'(\theta_n) = 0.$$

By way of contradiction, suppose otherwise. Then, there exists $\varepsilon > 0$ such that there exist infinite sequences $m_k, n_k, k = 1, 2, \dots$, such that

$$|f'(\theta_{m_k})| \leq \varepsilon, \quad |f'(\theta_{n_k})| \geq 2\varepsilon, \quad |f'(\theta_i)| \geq \varepsilon \quad (20)$$

for all $m_k + 1 \leq i \leq n_k$. It then follows that

$$\begin{aligned} \varepsilon &\leq |f'(\theta_{n_k}) - f'(\theta_{m_k})| \\ &= O(|\theta_{n_k} - \theta_{m_k}|) \\ &= O\left(\sum_{i=m_k}^{n_k-1} a_i |f'(\theta_i)|\right) + O\left(\left|\sum_{i=m_k}^{n_k-1} a_i R_i(\theta_i)\right|\right) \\ &= O\left(\sum_{i=m_k}^{n_k-1} a_i\right) + O\left(\left|\sum_{i=m_k}^{n_k-1} a_i R_i(\theta_i)\right|\right). \end{aligned} \quad (21)$$

As in the proof of Lemma 2, using the decomposition

$$\begin{aligned} R_n(\theta_n) &= g_{n^b}(\theta_n) - f'(\theta_n) \\ &= g_{n^b}(\theta_n) - f'_{n^b}(\theta_n) + f'_{n^b}(\theta_n) - f'(\theta_n), \end{aligned}$$

we deduce that $\sum_{n=0}^{\infty} a_n R_n(\theta_n)$ converges almost surely, and hence $\left|\sum_{i=m_k}^{n_k-1} a_i R_i(\theta_i)\right|$ tends to 0 as k goes to ∞ . On the other hand, by (20), we have

$$\varepsilon^2 \sum_{i=m_k}^{n_k-1} a_i \leq \sum_{i=m_k}^{\infty} a_i (f'(\theta_i))^2.$$

This implies that as k tends to ∞ , $\sum_{i=m_k}^{n_k-1} a_i$ tends to zero, which, together with (21), further implies that

$$\varepsilon \leq \lim_{k \rightarrow \infty} |f'(\theta_{n_k}) - f'(\theta_{m_k})| = 0,$$

a contradiction. \square

Remark 11: Apparently, the fact that $\{f(\theta_n)\}$ converges almost surely does not necessarily imply that $\{\theta_n\}$ converges almost surely. In the following sections, we will prove that, under some assumptions, $\{\theta_n\}$ does converge almost surely.

VI. SOME ESTIMATIONS

In this section, assuming (V.a), we will derive some estimations that will be used in the later sections for convergence analysis.

For any $j \in \mathbb{N}$, let

$$A_j = a_1 + a_2 + \dots + a_{j-1},$$

and for any $h > 0$ and any $n \in \mathbb{N}$, define

$$t(n, h) \triangleq \min\{k : a_n + a_{n+1} + \dots + a_{k-1} \geq h\}.$$

Now, for any fixed $n_0 \in \mathbb{N}$, recursively define

$$n_{k+1} = t(n_k, h).$$

One then verifies that for k sufficiently large,

$$A_{n_{k+1}} - A_{n_k} = \hat{O}(h), \quad n_k = \hat{O}(k^{1/(1-a)}), \quad (22)$$

where by $M = \hat{O}(N)$, we mean that there exist positive constants C_1, C_2 such that

$$C_1 N \leq M \leq C_2 N.$$

Now, an iterated application of

$$\theta_{n+1} - \theta_n = a_n f'(\theta_n) + a_n R_n(\theta_n)$$

yields

$$\begin{aligned} \theta_k &= \theta_n + \sum_{i=n}^{k-1} a_i f'(\theta_i) + \sum_{i=n}^{k-1} a_i R_i(\theta_i) \\ &= \theta_n + (A_k - A_n) f'(\theta_n) + \sum_{i=n}^{k-1} a_i R_i(\theta_i) \\ &\quad + \sum_{i=n}^{k-1} a_i (f'(\theta_i) - f'(\theta_n)) \\ &= \theta_n + R_{n,k}, \end{aligned}$$

where

$$R_{n,k} = \sum_{i=n}^{k-1} a_i R_i(\theta_i) + \sum_{i=n}^{k-1} a_i (f'(\theta_i) - f'(\theta_n)). \quad (23)$$

Similarly, an iterated application of

$$f(\theta_{n+1}) - f(\theta_n) = a_n f'^2(\theta_n) + \hat{R}_n(\theta_n)$$

yields

$$\begin{aligned} f(\theta_k) - f(\theta_n) &= \int_0^1 f'(\theta_n + t(\theta_k - \theta_n))(\theta_k - \theta_n) dt \\ &= \int_0^1 f'(\theta_n)(\theta_k - \theta_n) dt \\ &\quad + \int_0^1 (f'(\theta_n + t(\theta_k - \theta_n)) - f'(\theta_n))(\theta_k - \theta_n) dt \\ &= f'(\theta_n)((A_k - A_n) f'(\theta_n) + R_{n,k}) \\ &\quad + \int_0^1 (f'(\theta_n + t(\theta_k - \theta_n)) - f'(\theta_n))(\theta_k - \theta_n) dt \\ &= (A_k - A_n) f'^2(\theta_n) + f'(\theta_n) R_{n,k} \\ &\quad + \int_0^1 (f'(\theta_n + t(A_k - A_n)) - f'(\theta_n))(A_k - A_n) dt \\ &= (A_k - A_n) f'^2(\theta_n) + \hat{R}_{n,k}(\theta_n), \end{aligned} \quad (24)$$

(25)

where

$$\hat{R}_{n,k}(\theta_n) = f'(\theta_n)R_{n,k} + \int_0^1 (f'(\theta_n + t(A_k - A_n)) - f'(\theta_n))(A_k - A_n)dt. \quad (26)$$

The following lemma introduces a positive random variable, \tilde{C}_0 , and a constant, τ , which will be referred to throughout the rest of the paper.

Lemma 3: There exists a positive random variable \tilde{C}_0 such that for all n and for any $\tau > 0$ with $2a + b - 3b\beta - 2\tau > 1$,

$$\sup_{k \geq n} \left| \sum_{i=n}^k a_i R_i(\theta_i) \right| \leq \tilde{C}_0 n^{-\tau} \text{ a.s.}$$

Proof: For any $\tau > 0$ with $2a + b - 3b\beta - 2\tau > 1$, as in the proof of Lemma 2, we deduce that $\sum_{i=1}^{\infty} i^\tau a_i R_i$ converges almost surely. Letting

$$T_n \triangleq \sum_{i=1}^n i^\tau a_i R_i(\theta_i),$$

we then have for any $k \geq n$,

$$\begin{aligned} \sum_{i=n}^k a_i R_i(\theta_i) &= \sum_{i=n}^k (i^\tau a_i R_i(\theta_i)) i^{-\tau} \\ &= \sum_{i=n}^k (T_i - T_{i-1}) i^{-\tau} \\ &= \sum_{i=n}^k T_i i^{-\tau} - \sum_{i=n+1}^k T_{i-1} i^{-\tau} \\ &= \sum_{i=n}^k T_i i^{-\tau} - \sum_{i=n}^{k-1} T_i (i+1)^{-\tau} \\ &= T_k k^{-\tau} + \sum_{i=n}^{k-1} (i^{-\tau} - (i+1)^{-\tau}) T_i \\ &\leq (k^{-\tau} + \sum_{i=n}^{k-1} (i^{-\tau} - (i+1)^{-\tau})) \sup_i T_i \\ &= n^{-\tau} \sup_i T_i, \end{aligned}$$

which immediately implies the lemma. \square

In the following, to avoid notational cumbersomeness, we will use C to denote a positive constant, which may not carry the same value upon each of its appearances.

Lemma 4: Let $0 < h < 1$ and \tilde{C}_0, τ be as in Lemma 3, then we have

(1) there exists a constant $C > 0$ such that

$$|f'(\theta_{t(n,h)})| \leq C(\tilde{C}_0 n^{-\tau} + |f'(\theta_n)|).$$

(2) there exists a constant $C > 0$ such that

$$|\theta_{t(n,h)} - \theta_n| \leq C(\tilde{C}_0 n^{-\tau} + h|f'(\theta_n)|).$$

(3) there exists a constant $C > 0$ such that

$$|R_{n,t(n,h)}| \leq C(\tilde{C}_0 n^{-\tau} + h^2|f'(\theta_n)|).$$

(4) there exists a constant $C > 0$ such that

$$|\hat{R}_{n,t(n,h)}| \leq C(\tilde{C}_0^2 n^{-2\tau} + \tilde{C}_0 n^{-\tau} |f'(\theta_n)| + h^2 |f'(\theta_n)|^2).$$

(5) there exists a constant $C > 0$ such that

$$f(\theta_n) - f(\theta_{t(n,h)}) \leq -(3/4 - 3Ch/2)h|f'(\theta_n)|^2 + C\tilde{C}_0^2 n^{-2\tau} (1 + 1/(2h^2)).$$

(6) there exists $C > 0$ such that for sufficiently small h

$$2(f(\theta_n) - f(\theta_{t(n,h)})) + |f'(\theta_n)||\theta_{t(n,h)} - \theta_n| \leq (C + 1/(2h^2))\tilde{C}_0^2 n^{-2\tau}.$$

(7) for any $\tau' < \tau$, there exists a positive constant C such that for sufficiently small h , we have

$$|\theta_{t(n,h)} - \theta_n| \leq Cn^{\tau'} (f(\theta_{t(n,h)}) - f(\theta_n)) + C\tilde{C}_0^2 n^{-\tau'}.$$

Proof: In this proof, for notational simplicity, we will write $t(n, h)$ as k .

Note that there exists a positive constant C such that

$$\begin{aligned} |f'(\theta_k)| &\leq |f'(\theta_n)| + |f'(\theta_k) - f'(\theta_n)| \\ &\leq |f'(\theta_n)| + C|\theta_k - \theta_n| \\ &\leq |f'(\theta_n)| + C \sum_{i=n}^{k-1} a_i |f'(\theta_i)| + C \left| \sum_{i=n}^{k-1} a_i R_i(\theta_i) \right|, \end{aligned}$$

where we have applied (16). Applying Lemma 3, we then have

$$|f'(\theta_k)| \leq C\tilde{C}_0 n^{-\tau} + |f'(\theta_n)| + C \sum_{i=n}^{k-1} a_i |f'(\theta_i)|.$$

Applying Gronwall's lemma [22], we then have for n sufficiently large

$$\begin{aligned} |f'(\theta_k)| &\leq (C\tilde{C}_0 n^{-\tau} + |f'(\theta_n)|) \exp(C(a_n + a_{n+1} + \dots + a_{k-1})) \\ &\leq \exp(C)(C\tilde{C}_0 n^{-\tau} + |f'(\theta_n)|), \end{aligned}$$

where we have used the fact that for n large enough

$$a_n + a_{n+1} + \dots + a_{k-1} \approx h < 1.$$

We then have established (1).

It then follows from (1) that for some C

$$\begin{aligned} |\theta_k - \theta_n| &\leq \sum_{i=n}^{k-1} a_i |f'(\theta_i)| + \left| \sum_{i=n}^{k-1} a_i R_i(\theta_i) \right| \\ &\leq (A_k - A_n)(C\tilde{C}_0 n^{-\tau} + C|f'(\theta_n)|) + \tilde{C}_0 n^{-\tau}, \end{aligned}$$

which immediately implies (2).

Now, by (23) and (2), we have for some C

$$\begin{aligned} |R_{n,k}| &\leq \tilde{C}_0 n^{-\tau} + C \sum_{i=n}^{k-1} a_i |\theta_i - \theta_n| \\ &\leq \tilde{C}_0 n^{-\tau} + C^2 (A_k - A_n) (\tilde{C}_0 n^{-\tau} + (A_k - A_n) |f'(\theta_n)|), \end{aligned}$$

which establishes (3).

Furthermore, by (26), (2) and (3), we have

$$\begin{aligned} |\hat{R}_{n,k}| &\leq |f'(\theta_n)||R_{n,k}| + C|\theta_k - \theta_n|^2 \\ &\leq C\tilde{C}_0 n^{-\tau} |f'(\theta_n)| + C(A_k - A_n)^2 |f'(\theta_n)|^2 \\ &\quad + 2C^3 (\tilde{C}_0^2 n^{-2\tau} + (A_k - A_n)^2 |f'(\theta_n)|^2), \end{aligned}$$

which establishes (4).

It then follows from (24), (III) and (IV) and that for sufficiently large n

$$\begin{aligned}
f(\theta_n) - f(\theta_k) &\leq -(A_k - A_n)|f'(\theta_n)|^2 + |\hat{R}_{n,k}| \\
&\leq -3h/4|f'(\theta_n)|^2 + C(\tilde{C}_0^2 n^{-2\tau} \\
&\quad + \tilde{C}_0 n^{-\tau}|f'(\theta_n)| + h^2|f'(\theta_n)|^2) \\
&\leq -3h/4|f'(\theta_n)|^2 + C(\tilde{C}_0^2 n^{-2\tau} + \tilde{C}_0^2 n^{-2\tau}/(2h^2) \\
&\quad + h^2|f'(\theta_n)|^2/2 + h^2|f'(\theta_n)|^2) \\
&\leq -3h/4|f'(\theta_n)|^2 + C(\tilde{C}_0^2 n^{-2\tau}(1 + 1/(2h^2)) \\
&\quad + 3h^2/2|f'(\theta_n)|^2) \\
&\leq -(3/4 - 3Ch/2)h|f'(\theta_n)|^2 + C\tilde{C}_0^2 n^{-2c}(1 + 1/(2h^2)),
\end{aligned}$$

which establishes (5).

It follows from (24), (3) and (4) that

$$\begin{aligned}
f(\theta_k) - f(\theta_n) &= |f'(\theta_n)||A_k - A_n|f'(\theta_n)| + \hat{R}_{n,k} \\
&= |f'(\theta_n)||\theta_k - \theta_n + R_{n,k}| + \hat{R}_{n,k} \\
&\geq |f'(\theta_n)|(|\theta_k - \theta_n| - |R_{n,k}|) - \hat{R}_{n,k} \\
&\geq |f'(\theta_n)||\theta_k - \theta_n| - C(\tilde{C}_0^2 n^{-2\tau} + \tilde{C}_0 n^{-\tau}|f'(\theta_n)| \\
&\quad + (A_k - A_n)^2|f'(\theta_n)|^2),
\end{aligned}$$

which implies that

$$\begin{aligned}
f(\theta_n) - f(\theta_k) + |f'(\theta_n)||\theta_k - \theta_n| &\leq C(\tilde{C}_0^2 n^{-2\tau} + \tilde{C}_0 n^{-\tau}|f'(\theta_n)| + (A_k - A_n)^2|f'(\theta_n)|^2) \\
&\leq C(\tilde{C}_0^2 n^{-2\tau}(1 + 1/(2h^2)) + 3h^2/2|f'(\theta_n)|^2).
\end{aligned}$$

Applying (V), we then have for sufficiently small h ,

$$\begin{aligned}
f(\theta_n) - f(\theta_k) + |f'(\theta_n)||\theta_k - \theta_n| &\leq 2C(1 + 1/(2h^2))\tilde{C}_0^2 n^{-2\tau} + f(\theta_k - f(\theta_n)),
\end{aligned}$$

which can be rewritten as

$$\begin{aligned}
2(f(\theta_n) - f(\theta_k)) + |f'(\theta_n)||\theta_k - \theta_n| &\leq 2C(1 + 1/(2h^2))\tilde{C}_0^2 n^{-2\tau},
\end{aligned}$$

which establishes (6).

We next prove (7). If $|f'(\theta_n)| \leq n^{-\tau'}$, applying (II), we deduce that

$$|\theta_k - \theta_n| \leq C\tilde{C}_0 n^{-\tau'} + Ch^2 n^{-\tau'}. \quad (27)$$

It follows from (24) and (4) that

$$\begin{aligned}
|f(\theta_k) - f(\theta_n)| &\leq (A_k - A_n)f'^2(\theta_n) + |\hat{R}_{n,k}(\theta_n)| \\
&\leq (A_k - A_n)f'^2(\theta_n) + C(\tilde{C}_0^2 n^{-2\tau} + \tilde{C}_0 n^{-\tau}|f'(\theta_n)| \\
&\quad + h^2|f'(\theta_n)|^2) \\
&\leq (A_k - A_n)f'^2(\theta_n) + C(\tilde{C}_0^2 n^{-2\tau'} + \tilde{C}_0 n^{-\tau'}|f'(\theta_n)| \\
&\quad + h^2|f'(\theta_n)|^2) \\
&\leq C(\tilde{C}_0^2 n^{-2\tau'}(1 + 1/(2h^2))) + (h + 3Ch^2/2)|f'(\theta_n)|^2,
\end{aligned}$$

which, together with (27), immediately implies that for some C ,

$$\begin{aligned}
|\theta_k - \theta_n| &\leq n^{\tau'}(f(\theta_k) - f(\theta_n)) + n^{\tau'}|f(\theta_k) - f(\theta_n)| \\
&\quad + C\tilde{C}_0 n^{-\tau'} + Ch^2 n^{-\tau'} \\
&\leq n^{\tau'}(f(\theta_k) - f(\theta_n)) + C(\tilde{C}_0^2 n^{-2\tau'}(1 + 1/(2h^2))) \\
&\quad + (h + 3Ch^2/2)|f'(\theta_n)|^2 + C\tilde{C}_0 n^{-\tau'} + Ch^2 n^{-\tau'}.
\end{aligned} \quad (28)$$

On the other hand, if $|f'(\theta_n)| \geq n^{-\tau'}$, applying (6), we deduce that

$$\begin{aligned}
|\theta_k - \theta_n| &\leq 2|f'(\theta_n)|^{-1}(f(\theta_k) - f(\theta_n)) \\
&\quad + (C + 1/(2h^2))|f'(\theta_n)|^{-1}\tilde{C}_0^2 n^{-2\tau} \\
&\leq 2n^{\tau'}(f(\theta_k) - f(\theta_n)) + (C + 1/(2h^2))\tilde{C}_0^2 n^{-\tau'}.
\end{aligned} \quad (29)$$

Combining (28) and (29), we then have established (7). \square

VII. RATE OF CONVERGENCE WITH CONCAVITY

In this section, we assume that

(VII.a) $f(\theta)$ is strictly concave with respect to θ . More precisely, there exists $\hat{\epsilon} > 0$ such that for any $\theta_1, \theta_2 \in \Theta$,

$$f_t''(t\theta_1 + (1-t)\theta_2) \geq \hat{\epsilon},$$

for all $0 \leq t \leq 1$.

Here, let us note that (VII.a), together with Theorem 2, implies (V.a), which, together with Theorem 4, further implies that, with probability 1, θ_n converges to the unique global maximum θ^* as n tends to ∞ . With Assumption (VII.a), which, as argued in Section IX, can be satisfied for a class of finite-state channels, we will derive the convergence rate of $\{\theta_n\}$. Again, for notational convenience only, we assume that $\Theta = \mathbb{R}$ in the proofs.

From

$$\theta_{n+1} - \theta_n = a_n f'(\theta_n) + a_n R_n(\theta_n),$$

trivially we have

$$\Delta_{n+1} - \Delta_n = -a_n f'(\theta_n) - a_n R_n(\theta_n),$$

where

$$\Delta_n \triangleq (\theta^* - \theta_n).$$

It immediately from the above two conditions that for θ sufficiently close to θ^*

$$f(\theta) = \hat{O}(|\theta^* - \theta|^2), \quad f'(\theta) = \hat{O}(|\theta^* - \theta|). \quad (30)$$

So, if θ_n is sufficiently close to θ^* , we will have

$$f(\theta_n) = \hat{O}(\Delta_n^2), \quad f'(\theta_n) = \hat{O}(|\Delta_n|).$$

Throughout the paper, by $M = \tilde{O}(N)$, we mean that there exists a positive random variable \tilde{C} such that with probability 1,

$$|M| \leq \tilde{C}N.$$

In this section, we will prove that Δ_n is at most of that order $\tilde{O}(n^{-\tau})$.

We first prove the following lemma.

Lemma 5: There exists $l \in \mathbb{N}$ such that

$$\liminf_{n \rightarrow \infty} n^\tau |\Delta_n| \leq l\tilde{C}_0.$$

Proof: Suppose, by way of contradiction, that for any l ,

$$n^\tau |\Delta_n| \geq l\tilde{C}_0, \quad (31)$$

as long as n is sufficiently large. First, pick n_0 sufficiently large such that (31) is satisfied and then recursively define

$$n_{k+1} = t(n_k, h).$$

for some $0 < h < 1$. We then have, for any feasible k ,

$$\theta_{n_{k+1}} = \theta_{n_k} + (A_{n_{k+1}} - A_{n_k})f'(\theta_{n_k}) + R_{n_k, n_{k+1}}.$$

It then follows from Lemma 4 (3) and (30) that $R_{n_k, n_{k+1}}$ is dominated by $|f'(\theta_{n_k})|$ as long as l is chosen sufficiently large and h is chosen sufficiently small. Noticing that due to the concavity of f , Δ_n always has the same sign as $f'(\theta_n)$, then we have

$$|\Delta_{n_{k+1}}| \leq |\Delta_{n_k}| - h/2|\Delta_{n_k}| \leq |\Delta_{n_k}|e^{-h/2},$$

a recursive application of which would yield

$$\Delta_{n_k} \leq \Delta_{n_0} e^{-kh/2}.$$

It then follows that for any k

$$\Delta_{n_0} n_k^\tau e^{-kh/2} \geq n_k^\tau \Delta_{n_k} \geq l\tilde{C}_0.$$

This, together with the fact that (see (22))

$$n_k = \hat{O}(k^{1/(1-a)}),$$

as k tends to infinity, implies that

$$\tilde{C}_0 \leq 0,$$

which is a contradiction. \square

Theorem 5:

$$|\Delta_n| = \tilde{O}(n^{-\tau}).$$

Proof: It is enough to prove that there exists an integer l such that for all n sufficiently large,

$$n^\tau |\Delta_n| \leq l\tilde{C}_0.$$

By way of contradiction, suppose otherwise. Then, by Lemma 5, for any l and arbitrarily large N , we can find $k_0 > m_0 > N$ such that

$$\begin{aligned} m_0^\tau \Delta_{m_0} &\leq 2l\tilde{C}_0, & k_0^\tau \Delta_{k_0} &\geq 3l\tilde{C}_0, \\ \min_{m_0 < n \leq k_0} n^\tau \Delta_n &> 2l\tilde{C}_0, & \max_{m_0 \leq n < k_0} n^\tau \Delta_n &\leq 3l\tilde{C}_0. \end{aligned} \quad (32)$$

Now, for some $0 < h < 1$, let $m_1 = t(m_0, h)$. Then, for any $m_0 \leq n \leq m_1$, it follows from (32),

$$\theta_n - \theta_{m_0} = (A_n - A_{m_0})f'(\theta_{m_0}) + R_{m_0, n},$$

and

$$|R_{m_0, n}| \leq C(m_0^{-\tau}\tilde{C}_0 + (A_n - A_{m_0})^2|f'(\theta_{m_0})|)$$

$$|\Delta_n - \Delta_{m_0}| = O(m_0^{-\tau})\tilde{C}_0.$$

Applying (32), we then deduce that for sufficiently small h

$$\begin{aligned} |n^\tau \Delta_n - m_0^\tau \Delta_{m_0}| &\leq n^\tau |\Delta_n - \Delta_{m_0}| + (n^\tau - m_0^\tau)\Delta_{m_0} \\ &\leq O(m_0^\tau)O(m_0^{-\tau})\tilde{C}_0 + o(m_0^\tau)2lm_0^{-\tau}\tilde{C}_0, \end{aligned}$$

where we have used the fact that

$$n^\tau = O(m_0^\tau), \quad n^\tau - m_0^\tau = o(m_0^\tau).$$

It then follows that, with l large enough and h small enough, we have

$$|n^\tau \Delta_n - m_0^\tau \Delta_{m_0}| \leq l\tilde{C}_0.$$

In particular, we have

$$\begin{aligned} |(m_0 + 1)^\tau \Delta_{m_0+1} - m_0^\tau \Delta_{m_0}| &\leq l\tilde{C}_0 \text{ and } |m_1^\tau \Delta_{m_1} - m_0^\tau \Delta_{m_0}| \\ &\leq l\tilde{C}_0, \end{aligned}$$

which further implies that

$$m_0^\tau \Delta_{m_0} \geq l\tilde{C}_0 \quad \text{and} \quad m_1 < k_0,$$

respectively.

Now, for some $0 < h < 1$, we have

$$\theta_{m_1} = \theta_{m_0} + (A_{m_1} - A_{m_0})f'(\theta_{m_0}) + R_{m_0, m_1},$$

and

$$|R_{m_0, m_1}| \leq C(m_0^{-\tau}\tilde{C}_0 + (A_{m_1} - A_{m_0})^2|f'(\theta_{m_0})|).$$

As in the proof of Lemma 5, if l is chosen large enough, then $|f'(\theta_{m_0})|$ will dominate $|R_{m_0, m_1}|$. Again, due to the concavity of f (Assumption (VII.a)), Δ_{m_0} always has the same sign as $f'(\theta_{m_0})$, then for sufficiently small $h > 0$, we have

$$|\Delta_{m_1}| \leq |\Delta_{m_0}| - h/2|\Delta_{m_0}|.$$

\square Then, for m_0 sufficiently large such that

$$m_1^\tau < m_0^\tau/(1 - h/2),$$

we have

$$m_1^\tau |\Delta_{m_1}| \leq m_1^\tau |\Delta_{m_0}|(1 - h/2) < m_0^\tau |\Delta_{m_0}| \leq 2l\tilde{C}_0,$$

which is a contradiction to (32). \square

VIII. RATE OF CONVERGENCE WITHOUT CONCAVITY

Throughout this section, we will assume (V.a) and (VIII.a) with probability 1, $\theta_n \in \mathcal{Q}$ for all n , where \mathcal{Q} is a compact subset of Θ , and we will derive the rate of convergence of our algorithm under these two assumptions. Again, for notational convenience only, we assume that $\Theta = \mathbb{R}$.

As one of the main results in real algebraic geometry, the Lojasiewicz inequality [10], among many other applications, has been widely applied to the convergence analysis of a broad class of dynamical systems. In this section, we will first use the ‘‘function’’ version of the Lojasiewicz inequality (Lemma 6) to prove that $\{\theta_n\}$ converges almost surely and derive the convergence rate, which can be further used to

derive the convergence rate of $\{\theta_n\}$. Then, using the ‘‘variable’’ version of the Lojasiewicz inequality (Lemma 10), the rate of convergence can be refined. The above-mentioned framework is essentially due to Tadic [56], however, a comprehensive adaptation to our settings has been done in this section.

Following [56], we state the ‘‘function’’ version of the Lojasiewicz inequality as below.

Lemma 6: For any compact set $Q \subset \Theta$ and real number $z \in f(Q)$, there exist real numbers $\delta_{Q,z} \in (0, 1)$, $\mu_{Q,z} \in (1, 2]$ and $M_{Q,z} \in [1, \infty)$ such that

$$|f(\theta) - z| \leq M_{Q,z} |f'(\theta)|^{\mu_{Q,z}}$$

for all $\theta \in Q$ satisfying $|f(\theta) - z| \leq \delta_{Q,z}$.

From now on, we will set $\hat{f} = \lim_{n \rightarrow \infty} f(\theta_n)$ and write $\mu = \mu_{Q,\hat{f}}$. Define

$$\hat{\Delta}_n \triangleq \hat{f} - f(\theta_n).$$

We first prove the following lemma.

Lemma 7: There exists a positive integer l such that for all n sufficiently large,

$$n^{\mu\tau} \hat{\Delta}_n \geq -l\tilde{C}_0^\mu.$$

Proof: Suppose, by way of contradiction, that for any l , there exists some n_0 ,

$$n_0^{\mu\tau} \hat{\Delta}_{n_0} < -l\tilde{C}_0^\mu. \quad (33)$$

Then, by Lemma 4 (5), we have, for some $0 < h < 1$,

$$f(\theta_{n_0}) - f(\theta_{t(n_0,h)}) \leq -(3/4 - 3Ch/2)h|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/(2h^2)),$$

which implies for h sufficiently small,

$$\begin{aligned} \hat{\Delta}_{\theta_{t(n_0,h)}} - \hat{\Delta}_{n_0} &\leq -(3/4 - 3Ch/2)h|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/2h^2) \\ &\leq -h/2|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/(2h^2)). \end{aligned}$$

Choosing l sufficiently large, then by Lemma 6 and (33), we deduce that for n large enough,

$$-h/2|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} \leq -h/4|f'(\theta_{n_0})|^2,$$

and therefore

$$\hat{\Delta}_{t(n_0,h)} - \hat{\Delta}_{n_0} \leq -h/4|f'(\theta_{n_0})|^2. \quad (34)$$

We then have

$$\hat{\Delta}_{t(n_0,h)} \leq \hat{\Delta}_{n_0} \leq -l\tilde{C}_0^\mu n_0^{-\mu\tau} \leq -l\tilde{C}_0^\mu t(n_0, h)^{-\mu\tau}.$$

Henceforth, recursively define

$$n_{k+1} = t(n_k, h).$$

It then follows that for any k ,

$$\hat{\Delta}_{n_k} \leq \hat{\Delta}_{n_0} \leq -l\tilde{C}_0^\mu n_0^{-\mu\tau} < 0,$$

which is a contradiction to the fact that almost surely

$$\lim_{k \rightarrow \infty} \hat{\Delta}_{n_k} = 0.$$

□

In the remainder of this section, define

$$\hat{\tau} = \min(\mu\tau, \mu(1-a)/(2-\mu)).$$

Lemma 8: There exists a positive integer l such that

$$\liminf_{n \rightarrow \infty} n^{\hat{\tau}} \hat{\Delta}_n \leq l\tilde{C}_0^\mu$$

almost surely.

Proof: Suppose, by way of contradiction, that for any l , we have

$$n^{\hat{\tau}} \hat{\Delta}_n \geq l\tilde{C}_0^\mu, \quad (35)$$

for all n sufficiently large. By Lemma 4 (5), for any $0 < h < 1$, we have for n_0 large enough,

$$f(\theta_{n_0}) - f(\theta_{t(n_0,h)}) \leq -(3/4 - 3Ch/2)h|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/(2h^2)),$$

which implies that for h sufficiently small

$$\begin{aligned} \hat{\Delta}_{t(n_0,h)} - \hat{\Delta}_{n_0} &\leq -(3/4 - 3Ch/2)h|f'(\theta_{n_0})|^2 \\ &\quad + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/(2h^2)) \\ &\leq -h/2|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} (1 + 1/(2h^2)). \end{aligned}$$

Choosing l sufficiently large, then by Lemma 6 and (33), we deduce that sufficiently large n ,

$$-h/2|f'(\theta_{n_0})|^2 + C\tilde{C}_0^2 n_0^{-2\tau} \leq -h/4|f'(\theta_{n_0})|^2,$$

and therefore

$$\hat{\Delta}_{t(n_0,h)} - \hat{\Delta}_{n_0} \leq -h/4|f'(\theta_{n_0})|^2. \quad (36)$$

Now, recursively define

$$n_{k+1} = t(n_k, h).$$

An iterated application of (36) yields for some constant C_1 ,

$$\hat{\Delta}_{n_{k+1}} - \hat{\Delta}_{n_k} \leq -C_1 h \hat{\Delta}_{n_k}^{2/\mu}.$$

We then have two cases:

Case $\mu = 2$: For this case, we have,

$$\hat{\Delta}_{n_{k+1}} \leq (1 - C_1 h) \hat{\Delta}_{n_k}.$$

Recursively, we deduce that

$$\hat{\Delta}_{n_k} \leq \hat{\Delta}_{n_0} (1 - C_1 h)^k \leq \hat{\Delta}_{n_0} e^{-C_1 h k},$$

which implies that for any k ,

$$\hat{\Delta}_{n_0} n_k^{\mu\tau} e^{-C_1 h k} \geq n_k^{\mu\tau} \hat{\Delta}_{n_k} \geq l\tilde{C}_0^\mu.$$

This, however, will yield $\tilde{C}_0 \leq 0$ when we take k to ∞ , which is a contradiction.

Case $\mu < 2$: For this case, it follows from

$$\hat{\Delta}_{n_k} - \hat{\Delta}_{n_{k+1}} \geq C_1 h \hat{\Delta}_{n_k}^{2/\mu}.$$

that

$$\int_{\hat{\Delta}_{n_{k+1}}}^{\hat{\Delta}_{n_k}} \frac{1}{u^{2/\mu}} du \geq \int_{\hat{\Delta}_{n_{k+1}}}^{\hat{\Delta}_{n_k}} \frac{1}{\hat{\Delta}_{n_k}^{2/\mu}} du = \frac{\hat{\Delta}_{n_k} - \hat{\Delta}_{n_{k+1}}}{\hat{\Delta}_{n_k}^{2/\mu}} \geq Ch,$$

which implies that for some positive constant C_2

$$\hat{\Delta}_{n_{k+1}}^{-2/\mu+1} - \hat{\Delta}_{n_k}^{-2/\mu+1} \geq C_2 h.$$

Recursively, we deduce that

$$\hat{\Delta}_{n_k}^{-2/\mu+1} \geq \hat{\Delta}_{n_0}^{-2/\mu+1} + C_2 h k,$$

and furthermore

$$\begin{aligned} n_k^{\hat{\tau}(-2+\mu)/\mu} \hat{\Delta}_{n_k}^{(-2+\mu)/\mu} &\geq n_k^{\hat{\tau}(-2+\mu)/\mu} \hat{\Delta}_{n_0}^{-2/\mu+1} \\ &\quad + C_2 n_k^{\hat{\tau}(-2+\mu)/\mu} k h. \end{aligned}$$

It then follows from (35) and (22) that

$$\begin{aligned} \tilde{C}_0^{-2+\mu} l^{(-2+\mu)/\mu} &\geq n_k^{\hat{\tau}(-2+\mu)/\mu} \hat{\Delta}_{n_k}^{(-2+\mu)/\mu} \\ &\geq O(k n_k^{\hat{\tau}(-2+\mu)/\mu}) \\ &\geq O(k^{\hat{\tau}(-2+\mu)/(-a+1)\mu+1}). \end{aligned}$$

Now, one verifies that this gives us a contradiction if we take k, l to ∞ , as long as

$$\hat{\tau} \leq \mu(1-a)/(2-\mu),$$

or, equivalently,

$$\hat{\tau}(-2+\mu)/(-a+1)\mu+1 \geq 0.$$

□

Lemma 9: There exists an integer l such that for all n sufficiently large,

$$n^{\hat{\tau}} \Delta_n \leq l \tilde{C}_0^2.$$

Proof: By way of contradiction, suppose otherwise. Then, by Lemma 8, for any l and arbitrarily large N , we can find $k_0 > m_0 > N$ such that

$$\begin{aligned} m_0^{\hat{\tau}} \Delta_{m_0} &\leq 2l \tilde{C}_0^2, & k_0^{\hat{\tau}} \Delta_{k_0} &\geq 3l \tilde{C}_0^2, \\ \min_{m_0 < n \leq k_0} n^{\hat{\tau}} \Delta_n &> 2l \tilde{C}_0^2, & \max_{m_0 \leq n < k_0} n^{\hat{\tau}} \Delta_n &\leq 3l \tilde{C}_0^2. \end{aligned} \quad (37)$$

For some $0 < h < 1$, let $m_1 = t(m_0, h)$. For any $m_0 \leq n \leq m_1$, as in the proof of Theorem 7, we derive

$$\hat{\Delta}_n - \hat{\Delta}_{m_0} \leq -h/4 |f'(\theta_{m_0})|^2, \quad (38)$$

which, together with Theorem 7 and (37), implies that

$$f'(\theta_{m_0})^2 \leq 4/h \tilde{C}_0^2 O(m_0^{-\hat{\tau}}) + \tilde{C}_0^2 O(m_0^{-\mu\tau}),$$

which, together with (24), further implies that for some $C > 0$,

$$\begin{aligned} |\hat{\Delta}_n - \hat{\Delta}_{m_0}| &\leq C h |f'(\theta_{m_0})|^2 + C \tilde{C}_0^2 m_0^{-2\tau} \\ &\leq \tilde{C}_0^2 O(m_0^{-\hat{\tau}}) + \tilde{C}_0^2 O(m_0^{-\mu\tau}) + C \tilde{C}_0^2 m_0^{-2\tau}. \end{aligned}$$

It then follows that for sufficiently small h

$$\begin{aligned} |n^{\hat{\tau}} \hat{\Delta}_n - m_0^{\hat{\tau}} \hat{\Delta}_{m_0}| &\leq n^{\hat{\tau}} |\hat{\Delta}_n - \hat{\Delta}_{m_0}| + (n^{\hat{\tau}} - m_0^{\hat{\tau}}) \hat{\Delta}_{m_0} \\ &= O(m_0^{\hat{\tau}}) (\hat{\Delta}_n - \hat{\Delta}_{m_0}) + o(m_0^{\hat{\tau}}) \hat{\Delta}_{m_0} \\ &\leq l \tilde{C}_0^2, \end{aligned}$$

where we have used the fact that

$$n^{\hat{\tau}} = O(m_0^{\hat{\tau}}), \quad n^{\hat{\tau}} - m_0^{\hat{\tau}} = o(m_0^{\hat{\tau}}).$$

In particular, we have

$$|(m_0+1)^{\hat{\tau}} \hat{\Delta}_{m_0+1} - m_0^{\hat{\tau}} \hat{\Delta}_{m_0}| \leq l \tilde{C}_0^2,$$

and

$$|m_1^{\hat{\tau}} \hat{\Delta}_{m_1} - m_0^{\hat{\tau}} \hat{\Delta}_{m_0}| \leq l \tilde{C}_0^2,$$

which further implies that

$$m_0^{\hat{\tau}} \hat{\Delta}_{m_0} \geq l \tilde{C}_0^2 \text{ and } m_1 < k_0,$$

respectively.

Setting $n = m_1$ and rewriting (38), we have for some constant C_1 ,

$$\hat{\Delta}_{m_1} - \hat{\Delta}_{m_0} \leq -C_1 h \hat{\Delta}_{m_0}^{2/\mu}.$$

We then consider two cases:

Case $\mu = 2$: For this case, we have for some positive constant C_1 ,

$$\hat{\Delta}_{m_1} \leq (1 - C_1 h) \hat{\Delta}_{m_0}.$$

Then for m_0 large enough,

$$\begin{aligned} m_1^{\hat{\tau}} \hat{\Delta}_{m_1} &\leq (1 - C_1 h) m_1^{\hat{\tau}} \hat{\Delta}_{m_0} \\ &= (1 - C_1 h) m_0^{\hat{\tau}} (1 + o(1)) \hat{\Delta}_{m_0} \\ &\leq 2l \tilde{C}_0^2, \end{aligned}$$

which yields a contradiction.

Case $\mu < 2$: For this case, as in the proof of Lemma 8, we have for some positive constant C_2 ,

$$\hat{\Delta}_{m_1}^{-2/\mu+1} \geq \hat{\Delta}_{m_0}^{-2/\mu+1} + C_2 h.$$

It then follows from (37) and (22) that for l sufficiently large

$$\begin{aligned} \hat{\Delta}_{m_1}^{(-2+\mu)/\mu} &\geq (2l \tilde{C}_0^2)^{(-2+\mu)/\mu} m_0^{-a+1} + C_2 h \\ &\geq (2l \tilde{C}_0^2)^{(-2+\mu)/\mu} m_1^{-a+1}, \end{aligned}$$

which implies that

$$m_1^{\hat{\tau}} \hat{\Delta}_{m_1} \leq 2l \tilde{C}_0^2,$$

a contradiction. □

The following theorem characterizes the rate of convergence of $\{f(\theta_n)\}$.

Theorem 6: With probability 1, we have

$$|\hat{\Delta}_n| = \tilde{O}(n^{-\hat{\tau}}).$$

Proof: It immediately follows from Lemmas 8 and 9. □

In the rest of this section, besides (V.a) and (VIII.a), we further assume (VIII.b) $\mu\tau \geq (1-a)$, and we will prove $\{\theta_n\}$ converges almost surely. Here, let us note that (VIII.b) can always be satisfied if a, b, β are appropriately chosen such that τ is sufficiently large.

The following theorem proves that when further assuming (VIII.b), $\{\theta_n\}$ converges almost surely, and moreover, it characterizes the rate of convergence of $\{\theta_n\}$.

Theorem 7: Further assuming (VIII.b), we have

$$\sup_{k \geq n} |\theta_k - \theta_n| = \tilde{O}(n^{-(\hat{\tau} - (1-a)/2)}).$$

Proof: In this proof, we set

$$\tau' = (\hat{\tau} + (1-a))/2.$$

For some $0 < h < 1$, starting from a fixed n_0 , recursively define

$$n_{k+1} = t(n_k, h).$$

Then, to prove the theorem, it suffices to prove that

$$\sup_{k \geq m} |\theta_{n_k} - \theta_{n_m}| = \tilde{O}(n_m^{-(\hat{\tau} + (1-a)/2)}). \quad (39)$$

Now, applying Lemma 4 (7), we deduce that for some $C > 0$

$$|\theta_{n_{i+1}} - \theta_{n_i}| \leq C n_i^{\tau'} (f(\theta_{n_{i+1}}) - f(\theta_{n_i})) + C \tilde{C}_0^2 n_i^{-\tau'}.$$

It then follows that for any $m \leq k$,

$$\begin{aligned} |\theta_{n_k} - \theta_{n_m}| &\leq \sum_{i=m}^{k-1} |\theta_{n_{i+1}} - \theta_{n_i}| \\ &\leq C \tilde{C}_0^2 \sum_{i=m}^{k-1} n_i^{-\tau'} + C \sum_{i=m}^{k-1} (u(\theta_{n_i}) - u(\theta_{n_{i+1}})) n_i^{\tau'} \\ &\leq C \tilde{C}_0^2 \sum_{i=m}^{k-1} n_i^{-\tau'} + C \sum_{i=m+1}^k (n_i^{\tau'} - n_{i-1}^{\tau'}) |u(\theta_{n_i})| \\ &\quad + C n_m^{\tau'} |u(\theta_{n_m})| + C n_k^{\tau'} |u(\theta_{n_k})|. \end{aligned}$$

Applying (22), we deduce that

$$\begin{aligned} \sum_{i=m}^{k-1} n_i^{-\tau'} &= \sum_{i=m}^{k-1} O(i^{-\tau'/(1-a)}) \\ &= O(m^{-\tau'/(1-a)+1}), \end{aligned}$$

and

$$\begin{aligned} \sum_{i=m+1}^k (n_i^{\tau'} - n_{i-1}^{\tau'}) |u(\theta_{n_i})| &= \sum_{i=m}^k O((i-1)^{y/(1-a)-1} i^{-\hat{\tau}/(1-a)}) \\ &= O(m^{y/(1-a)-\hat{\tau}/(1-a)}), \end{aligned}$$

and

$$\begin{aligned} n_m^{\tau'} |u(\theta_{n_m})| &= O(m^{\tau'/(1-a)} m^{-\hat{\tau}/(1-a)}) \\ &= O(m^{(\tau'-\hat{\tau})/(1-a)}), \end{aligned}$$

and

$$\begin{aligned} n_k^{\tau'} |u(\theta_{n_k})| &= O(k^{\tau'/(1-a)} k^{-\hat{\tau}/(1-a)}) \\ &= O(k^{(\tau'-\hat{\tau})/(1-a)}). \end{aligned}$$

We then deduce that

$$|\theta_{n_k} - \theta_{n_m}| = O(n_m^{(\tau'-\hat{\tau})}),$$

which immediately implies (39). \square

The following ‘‘variable’’ version of the Lojasiewicz inequality will be used to refine the rates of convergence of $\{\theta_n\}$ and $\{f(\theta_n)\}$.

Lemma 10: For each $\theta \in \Theta$, there exist real numbers $\delta_\theta \in (0, 1)$, $\mu_\theta \in (1, 2]$, $M_\theta \in [1, \infty)$ such that

$$|f(\theta') - f(\theta)| \leq M_\theta \|f'(\theta')\|^{\mu_\theta}$$

for all $\theta' \in \Theta$ satisfying $\|\theta' - \theta\| \leq \delta_\theta$.

From now on, let $\hat{\theta} = \lim_{n \rightarrow \infty} \theta_n$ and set $\mu = \mu_{\hat{\theta}}$. Then, with this redefined μ , going through exactly the same arguments as in the proof of Theorems 6 and 7, we have the following two theorems.

Theorem 8: For the above redefined μ , Theorems 6 holds, that is, with probability 1,

$$|\hat{\Delta}_n| = \tilde{O}(n^{-\hat{\tau}}).$$

Theorem 9: Further assuming (VIII.b) for the above redefined μ , Theorem 7 holds, that is,

$$|\theta_n - \hat{\theta}| = \tilde{O}(n^{-(\hat{\tau} - (1-a)/2)}).$$

IX. CAPACITY-ACHIEVING DISTRIBUTIONS OF A SPECIAL CLASS OF CHANNELS

In this section, we restrict our attention to a special class of input-restricted finite-state channels with certain parameterization and we prove that for such channels operating at high SNR regime, the capacity will only be achieved at the interior of the parameter space and our algorithm converges almost surely.

More specifically, recalling that X, Y denote the input, output processes of the channel over finite alphabets \mathcal{X}, \mathcal{Y} , respectively, we consider a class of parameterized memoryless channels such that

- (IX.a) the channel only has one state; in other words, at any time slot, the channel is characterized by the conditional probability $p(y|x)$.
- (IX.b) for some mixing finite-type constraint $F \subset \mathcal{X}^2$, $X \in \Pi_F$.
- (IX.c) the channel is parameterized by $\varepsilon \geq 0$ such that for each x and y , $p(y|x)(\varepsilon)$ is an analytic function of $\varepsilon \geq 0$, which is not identically 0.
- (IX.d) there is a one-to-one (not necessarily onto) mapping $\Phi : \mathcal{X} \rightarrow \mathcal{Y}$, such that for any $x \in \mathcal{X}$, $p(\Phi(x)|x)(0) = 1$.
- (IX.e) X is parameterized as in [60], that is,

$$\theta = (p(X_1 = w_1, X_2 = w_2) : (w_1, w_2) \notin F).$$

Under the above assumptions, ε can be regarded as a parameter that quantifies noise, and $\Phi(x)$ is the noiseless output corresponding to input x . The regime of ‘‘small ε ’’ corresponds to high SNR. Note that the output process $Y = Y(X, \varepsilon)$ depends on the input process X and the parameter value ε ; we will often suppress the notational dependence on ε or X , when it is clear from the context. Prominent examples of such families include input-restricted versions of the binary symmetric channel with crossover probability ε , and the binary erasure channel with erasure probability ε .

General SNR Regime: By using an asymptotic formula of $I(X; Y)$, we show that for the above-mentioned channels, the capacity-achieving X must be primitive.

Assume that X has period e with period classes D_1, D_2, \dots, D_e . Then, by the classical Perron-Frobenius theory, after necessary reindexing, its transition probability

matrix Π can be written as

$$\begin{matrix} & D_1 & D_2 & D_3 & \cdots & D_e \\ \begin{matrix} D_1 \\ D_2 \\ \vdots \\ D_{e-1} \\ D_e \end{matrix} & \begin{pmatrix} 0 & B_1 & 0 & \cdots & 0 \\ 0 & 0 & B_2 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & B_{e-1} \\ B_e & 0 & 0 & \cdots & 0 \end{pmatrix}, & (40) \end{matrix}$$

where we used the period classes to index the sub-blocks. In the following, let \mathbf{B} denote the set of all entry indices of Π corresponding to some B_k , that is,

$$\mathbf{B} = \{(i, j) : i \in D_k, j \in D_{k+1}, \text{ for } k = 1, \dots, e-1\} \\ \cup \{(i, j) : i \in D_e, j \in D_1\}.$$

Now, consider an analytic perturbation $\Pi(\delta)$ of Π , $\delta \geq 0$, where

- (IX.f) $\Pi(0) = \Pi$;
- (IX.g) for some $(i, j) \in \mathbf{B}$, $\Pi_{ij}(\delta)$ is not identically 0;
- (IX.h) for any $\delta \geq 0$, $\Pi(\delta)$ is still a stochastic matrix.

In other words, some non- B -entries in Π are analytically perturbed; as a result, Y is perturbed from $Y(0)$ to $Y(\delta)$. The following theorem describes the asymptotic behavior of $H(Y)$ under such a perturbation.

Theorem 10: Under the aboved-mentioned perturbation as in (IX.f)-(IX.h), there exist $C_1, C_2 > 0$ such that

$$C_1 \delta \log 1/\delta \leq H(Y(\delta)) - H(Y(0)) \leq C_2 \delta^{1/2}.$$

Proof: The proof is postponed to Appendix. \square

Remark 12: It follows from Condition (IX.a) that $H(Y|X)$ is linear with respect to \vec{p} . Theorem 10, together with this fact, implies that there exist C_1, C_2 such that

$$C_1 \delta \log 1/\delta \leq I(X(\delta); Y(\delta)) - I(X(0); Y(0)) \leq C_2 \delta^{1/2},$$

which implies that, for any irreducible but not primitive X , any perturbation of X as in (IX.f)-(IX.h) will strictly increase the mutual information. So, we conclude that the capacity-achieving X must be primitive, and thus for the purpose of computing the capacity, we can strengthen Condition (IX.b) to Condition (II.a), that is, $X \in \Pi_{F, \epsilon}$ for some $\epsilon > 0$.

High SNR Regime: At the high SNR regime, that is, when ϵ is close to 0, it has been established in [29] that there exists $\hat{\epsilon} > 0$ such that

- (IX.i) $I(X; Y)$ is strictly concave with respect to $\theta \in \Theta$, where Θ corresponds to $\Pi_{F, \hat{\epsilon}}$
- (IX.j) the capacity of the channel can be uniquely achieved within $\Pi_{F, \hat{\epsilon}}$.

As a consequence, we have the following theorem.

Theorem 11: For the channel as in (IX.a)-(IX.d) operating at the high SNR regime and sufficiently small $\hat{\epsilon}$, under the iteration in (3) with Θ set to correspond to $\Pi_{F, \hat{\epsilon}}$, $\{\theta_n\}$ converges almost surely and its limit corresponds to the capacity-achieving distribution.

Proof: Note that Condition (IX.i) and Theorem 4 imply Conditions (VII.a), which further implies that $\{\theta_n\}$ converges to a limit θ^* ; Condition (IX.j) implies that θ^* indeed corresponds to the capacity-achieving distribution. The theorem then immediately follows. \square

Example 1: Consider the binary symmetric channel with crossover probability $\epsilon > 0$. Let X be a binary input Markov chain with the transition probability matrix

$$\begin{bmatrix} 1 - \pi & \pi \\ 1 & 0 \end{bmatrix}, \quad (41)$$

where $0 \leq \pi \leq 1$. Apparently, X is supported on the so-called $(1, \infty)$ -RLL constraint [40], which simply means that the string “11” is forbidden. Let Y denote the corresponding output process. Assume that X is parameterized by $\theta = (p(00), p(01), p(10))$, where $p(10) = 1$ is in fact a constant. It can be checked that Conditions (IX.a)-(IX.d) are all satisfied, so when ϵ is sufficiently small, Conditions (IX.i)-(IX.j) are satisfied and thus Theorem 11 holds.

On the other hand, it has been shown that for the output process Y , as $\epsilon \rightarrow 0$,

$$H(Y) = H(X) + \frac{\pi(2 - \pi)}{1 + \pi} \epsilon \log(1/\epsilon) + O(\epsilon), \quad (42)$$

where the $O(\epsilon)$ -term is analytic with respect to p (see [29, Th. 2.18]). It then follows that

$$\begin{aligned} H(X|Y) &= H(X) + H(Y|X) - H(Y) \\ &= H(\epsilon) - \frac{\pi(2 - \pi)}{1 + \pi} \epsilon \log(1/\epsilon) + O(\epsilon), \end{aligned}$$

where $H(\epsilon) = \epsilon \log 1/\epsilon + (1 - \epsilon) \log 1/(1 - \epsilon)$. One can readily verify that $-\pi(2 - \pi)/(1 + \pi)$ is strictly convex with respect to θ , which implies the strict convexity (rather than concavity) of $H(X|Y)$ when ϵ is small enough. So, the concavity conjecture in [60] is not true in general, and thus the conditions guaranteeing the convergence of the GBAA are not satisfied.

X. NUMERICAL SIMULATIONS

Extensive numerical simulations of the algorithm have been conducted for a number of practical channels, (typically) fast convergence of $\{f(\theta_n)\}$ and $\{\theta_n\}$ have been observed.

Example 2: Consider a binary symmetric channel with crossover probability $\epsilon > 0$ as in Example 1, denote by $\text{BSC}(\epsilon)$. Let X be a binary input Markov chain supported on the so-called $(1, \infty)$ -RLL constraint [40] and assume the transition probability matrix of X takes the following form:

$$\begin{bmatrix} 1 - \pi & \pi \\ 1 & 0 \end{bmatrix}, \quad (43)$$

where $0 \leq \pi \leq 1$. The capacity of such a channel has been computed using the GBAA in [60]. For any fixed $\epsilon > 0$, setting $\theta = \pi$, our algorithm is observed to yield a sequence $\{\theta_n\}$ convergent to a limit θ^* . As shown in Figure 1, computations employing our algorithm yield essentially the same curve of the channel capacity C (as a function of ϵ) as the GBAA; the curve of θ^* is also plotted against ϵ in Figure 1.

Example 3: Consider a binary erasure channel with erasure probability $\epsilon > 0$, denoted by $\text{BEC}(\epsilon)$. Again, let X be a binary input Markov chain supported on the $(1, \infty)$ -RLL constraint and assume the transition probability matrix of X takes the

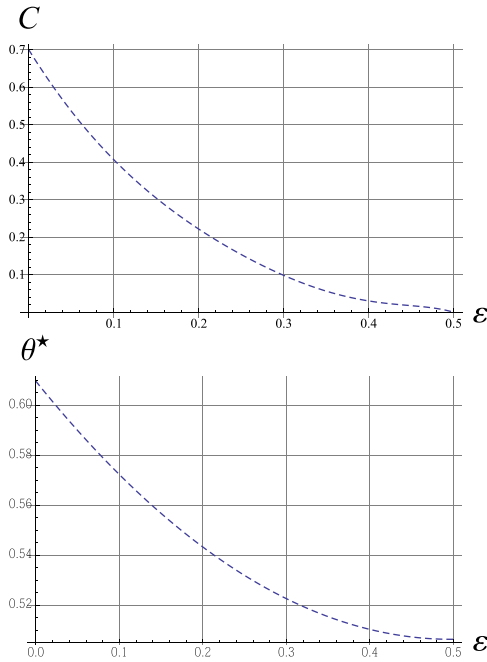


Fig. 1. BSC(ϵ): capacity C and capacity-achieving π versus ϵ .

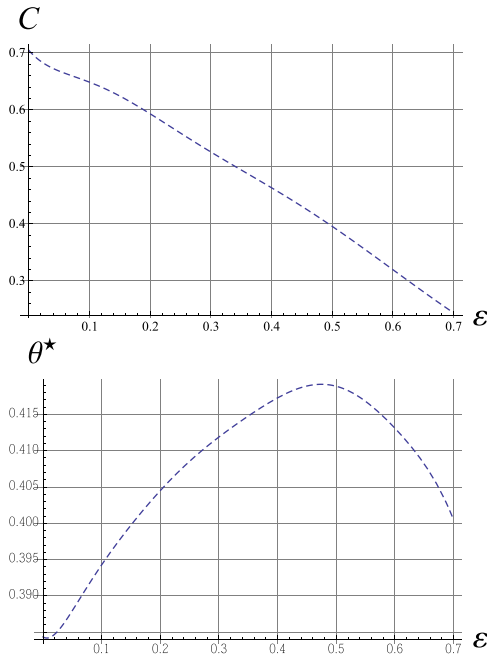


Fig. 2. BEC(ϵ): capacity C and capacity-achieving π versus ϵ .

same form as in (43). Such a channel has been examined in [38] and the authors show that $I(X; Y)$ is concave with respect to π , which implies that, when setting $\theta = \pi$, our algorithm yields an almost surely convergent sequence $\{\theta_n\}$ with a limit θ^* , which corresponds to the capacity achieving distribution. Indeed, the convergence of our algorithm has been observed [38], and the channel capacity C and the limit θ^* are plotted as functions of ϵ in Figure 2.

Example 4: The classical Gilbert-Elliott channel is a special kind of finite-state channel with two channel states, which,

independent from the binary channel inputs, will evolve according to a Markovian law. For the Gilbert-Elliott channel without any input constraint, it has been shown [44] that the capacity-achieving input is an i.i.d. *Bernoulli*(1/2) process. For testing purpose, we have implemented our algorithm for such Gilbert-Elliott channels and fast convergence of our algorithm to the capacity and its achieving distribution has been observed.

Example 5: The classical trapdoor channel was introduced by Blackwell [13] as a simple two-state channel whose behavior can be described as follows: At time 0, the channel starts with a ball, labeled “0” or “1”. Then, at each subsequent time slot: 1) a ball labeled “0” or “1” is inserted into the channel by the transmitter; 2) and the receiver receives one of the two balls in the channel with equal probability; 3) and the ball that does not exit the channel remains inside for the next channel use.

Apparently the above-mentioned trapdoor channel can be interpreted as a finite-state channel as follows: At each time slot, the state of the trapdoor channel is the label of the ball that is in the channel before the transmitter inserts a new ball. At time t , letting $S_{t-1} \in \{0, 1\}$ be the channel state, and $X_t \in \{0, 1\}$ be channel input (the label of the ball that is inserted) and Y_t be the corresponding output (the label of the ball that is received), the channel then can be characterized by the transition probabilities as tabulated below:

x_t	s_{t-1}	$p(y_t = 0 x_t, s_{t-1})$	$p(y_t = 1 x_t, s_{t-1})$
0	0	1	0
0	1	0.5	0.5
1	0	0.5	0.5
1	1	0	1

Computations employing our algorithm yields 0.4100 as a local maximum of $I(X; Y)$ over all binary Markov chains X with states “0” and “1”, and 0.5000 as a local maximum over all Markov chains X with states “00”, “01”, “10” and “11”. Here, we remark that when the Markovian restriction on the channel input is lifted, the capacity of the trapdoor channel is lower bounded [1] by 0.5 bits per channel use, and upper bounded [41] by 0.6610 bits per channel use.

Example 6: For any $0 \leq \alpha \leq 1$, the so-called POST(α) channels were first introduced in [61], which are simple finite-state channels where the state of the channel is the previous output. More precisely, the binary inputs $\{X_i\}_{i \geq 1}$ and the binary outputs $\{Y_i\}_{i \geq 1}$ of the POST(α) channel are related in the following way: if $X_i = Y_{i-1}$, then $Y_i = X_i$, else $Y_i = X_i \oplus Z_i$, where $Z_i \sim \text{Bernoulli}(\alpha)$. It has been shown that for such a channel, the feedback does not increase the capacity, which can be explicitly computed as $\log_2(1 + (1 - \alpha)\alpha^{\frac{\alpha}{1-\alpha}})$. As shown in Figure 3, when plotted against α , the capacity curve by our algorithm essentially coincides with the real capacity curve, and fast convergence of our algorithm has been observed for such channels.

Remark 13: While fast convergence of our algorithm has been typically observed in most of channels we have

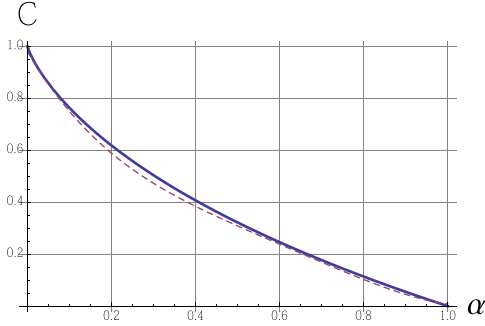


Fig. 3. POST(α) Channel: capacity C versus α (the solid curve represents the real capacity, while the dashed one is by our algorithm).

examined, there are some issues for the algorithm and its implementations:

- When the strict concavity of $I(X; Y)$ is absent, the algorithm may deliver a rather slow convergence. For instance, for the binary symmetric channel as in Example 2 operating at the low SNR regime (meaning ε is close to 0.5), compared to the high SNR regime (meaning ε is close to 0), the behavior of our algorithm is somewhat erratic and very low convergence has been observed.
- While the rate of convergence of our algorithm can be proven to be independent of choice of parameters, the speed of convergence in practice is affected by choices of the algorithm parameters, such as the starting point, θ_0 , the step sizes, a_n , and other parameters, a, b, α, β .
- Our algorithm is randomized in nature and its convergence is in the almost sure sense. So, the “capacity” computed using the algorithm can only be, at best, interpreted as “close” to the real capacity with high probability, which, will be of discounted significance when a definite value (or bounds) of the real capacity is needed.
- To compute the capacity of a finite-state channel with finite-type input constraints but without the Markovian assumption on the channel input, one can concatenate our algorithms implemented with Markov inputs of increasing orders. It is not surprising, however, that for some channels, the capacity achieving distribution can only be approximated by Markov chains with prohibitively large order, which may render the implementation/execution of our algorithm infeasible. In this regard, we have observed that our algorithm implemented for the trapdoor channel as in Example 5 with 4-state Markov chains converges considerably slower than that with 2-state Markov chains.

APPENDIX PROOF OF THEOREM 10

First of all, we define

$$\begin{aligned}
 & Z(\delta) \\
 &= Z(X_1^n(\delta)) \\
 &= \begin{cases} 0 & (X_i(\delta), X_{i+1}(\delta)) \in \mathbf{B} \text{ for all } i \in \{1, \dots, n-1\} \\ 1 & (X_i(\delta), X_{i+1}(\delta)) \notin \mathbf{B} \text{ for exactly one } i \in \{1, \dots, n-1\} \\ 2 & (X_i(\delta), X_{i+1}(\delta)) \notin \mathbf{B} \text{ for more than one } \\ & \quad i \in \{1, \dots, n-1\}. \end{cases}
 \end{aligned}$$

Next, applying the Birch bound [11], we derive the following key inequality for this proof:

$$\begin{aligned}
 & \frac{H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), Z(\delta))}{n-m} \\
 & \leq H(Y) \\
 & \leq \frac{H(Y_1^n|X_0(\delta), Z(\delta))}{n} + \frac{H(Z(\delta))}{n} + \frac{H(X_0(\delta))}{n}, \quad (44)
 \end{aligned}$$

for any $m \leq n$.

The Lower Bound Part: We first prove that there exists $C_1 > 0$ such that

$$H(Y(\delta)) \geq H(Y(0)) + C_1 \delta \log 1/\delta,$$

which immediately implies the lower bound part of the theorem. In this part, we set

$$n = \sqrt{\log \delta} \quad \text{and} \quad m = n/2. \quad (45)$$

By definition, we have

$$\begin{aligned}
 & \frac{H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), Z(\delta))}{n-m} \\
 &= \sum_{x_0} p^\delta(x_0, Z=0) H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), \\
 & \quad Z(\delta)=0)/(n-m) \\
 &+ \sum_{x_0} p^\delta(x_0, Z=1) H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), \\
 & \quad Z(\delta)=1)/(n-m) \\
 &+ \sum_{x_0} p^\delta(x_0, Z=2) H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), \\
 & \quad Z(\delta)=2)/(n-m). \\
 & \triangleq T_1 + T_2 + T_3
 \end{aligned}$$

where $p^\delta(x_0, Z=0)$ means $P(X_0(\delta) = x_0, Z(\delta) = 0)$.

We next give estimates for the each of three terms defined as above.

For T_3 , notice that $n\delta < 1$ for sufficiently small δ and then

$$\begin{aligned}
 \sum_{x_0} p^\delta(x_0, Z=2) & \leq n^2(C_0\delta)^2 + n^3(C_0\delta)^3 + \dots \\
 & \leq \frac{C_0^2}{1-nC_0\delta} n^2 \delta^2,
 \end{aligned}$$

for some $C_0 > 0$. It then follows that

$$\begin{aligned}
 T_3 &= \sum_{x_0} p^\delta(x_0, Z=2) \frac{H(Y_{m+1}^n(\delta)|Y_1^m(\delta), X_0(\delta), Z(\delta)=2)}{n-m} \\
 &\leq \sum_{x_0} p^\delta(x_0, Z=2) \frac{H(Y_{m+1}^n(\delta))}{n-m} \\
 &\leq \sum_{x_0} p^\delta(x_0, Z=2) \log |\mathcal{Y}| \\
 &= O(n^2 \delta^2). \quad (46)
 \end{aligned}$$

For T_2 , one verifies that for any x_0 , there exist constants $C_1, C_2 > 0$, $0 < \lambda_1 < \lambda_2 < 1$ such that

$$C_1 n \delta \lambda_1^n \leq p^\delta(y_1^n | x_0, Z=1) \leq C_2 n \delta \lambda_2^n.$$

Similarly, for any x_0 , there exist $C_3, C_4 > 0$, and the same $0 < \lambda_1 < \lambda_2 < 1$ as above such that

$$C_3 m \lambda_1^m \leq p^\delta(y_1^m | x_0, Z=1) \leq C_4 m \lambda_2^m.$$

It then follows that for any x_0 ,

$$C_5 \delta \lambda_2^n / \lambda_1^m \leq p^\delta(y_{m+1}^n | y_1^m, X_0, Z = 1) \leq C_6 \delta \lambda_2^n / \lambda_1^m,$$

which, together with (45), implies that

$$\begin{aligned} H(Y_{m+1}^n(\delta) | Y_1^m(\delta), X_0(\delta), Z(\delta) = 1) \\ = \hat{O}(\log 1/\delta) + O(n \log \lambda_2) + O(m \log \lambda_1). \end{aligned}$$

This, together with the fact

$$p(x_0, Z = 1) = \hat{O}(n\delta),$$

implies that

$$T_2 = \hat{O}(\delta \log 1/\delta) + O(n\delta \log \lambda_2) + O(m\delta \log \lambda_1). \quad (47)$$

For T_1 , notice that it can be rewritten as

$$T_1 = \frac{\sum p^\delta(y_1^n, x_0, Z = 0) \log p^\delta(y_{m+1}^n | y_1^m, x_0, Z = 0)}{n - m}.$$

One then verifies that

$$\begin{aligned} \left| p^\delta(y_1^n, x_0, Z = 0) - p^0(y_1^n, x_0, Z = 0) \right|_{\delta=0} \\ = O(n\delta) p^0(y_1^n, x_0, Z = 0), \end{aligned}$$

which implies that

$$\left| \frac{\sum p^0(y_1^n, x_0, Z = 0) \log p^\delta(y_{m+1}^n | y_1^m, x_0, Z = 0)}{n - m} - \frac{\sum p^\delta(y_1^n, x_0, Z = 0) \log p^\delta(y_{m+1}^n | y_1^m, x_0, Z = 0)}{n - m} \right| = O(n\delta).$$

When fixing x_0 and assuming $Z = 0$, the analyticity argument in [25] can be used to prove that

$$\sum p^0(y_1^n, x_0, Z = 0) \log p^\delta(y_{m+1}^n | y_1^m, x_0, Z = 0) / (n - m)$$

exponentially converges to an analytic function of δ . It then follows that for some $0 < \rho < 1$

$$T_1 = H(Y(0)) + O(\rho^m) + O(\delta). \quad (48)$$

Combining (46), (47) and (48), we then have

$$\begin{aligned} H(Y_{m+1}^n(\delta) | Y_1^m, X_0(\delta), Z(\delta)) / (n - m) \\ = H(Y(0)) + \hat{O}(\delta \log 1/\delta). \end{aligned}$$

The Upper Bound Part: We then prove that there exists $C_2 > 0$,

$$H(Y(\delta)) \leq H(Y(0)) + C_2 \delta^{1/2},$$

which immediately implies the upper bound part of the theorem. For this part, setting

$$n = \delta^{-1/2} \quad \text{and} \quad m = 0. \quad (49)$$

Using a parallel argument as in the lower bound part, we can still derive (46), (47) and (48) and then

$$H(Y_1^n | X_0(\delta), Z(\delta)) / n = H(Y)_{\delta=0,0} + O(\delta^{1/2}).$$

It can be verified that

$$p(Z(\delta) = 1) = O(n\delta), \quad p(Z(\delta) = 2) = O(n^2 \delta^2),$$

which, together with the straightforward observation that $H(X_0(\delta)) / n = O(1/n)$, implies that

$$\begin{aligned} H(Z(\delta)) &= - \sum_{i=0}^2 p(Z(\delta) = i) \log p(Z(\delta) = i) \\ &= O(n\delta \log \delta) + O(n\delta \log n), \end{aligned}$$

and consequently

$$\frac{H(Z(\delta))}{n} = O(\delta \log \delta) + O(\delta \log n).$$

The upper bound part then follows from all the above estimates and (44).

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