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RETAIL RENT WITH RESPECT TO DISTANCE FROM LIGHT RAIL TRANSIT STATIONS IN DALLAS AND DENVER

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47 Abstract

- 48 A growing body of recent research is challenging the assumptions underlying the half-mile-circle
- 49 in planning for development around transit stations. In this article we review this literature and
- 50 extend it to include retail land uses. We estimate the rent premium conferred on retail properties
- 51 in metropolitan Dallas and metropolitan Denver, both of which have extensive light rail transit
- 52 systems. We find that consistent with half-mile-circle assumptions, retail rent premiums extend
- only to about 0.30 mile from transit stations with half the premium dissipating after a few
- 54 hundred feet and three quarters within the first 0.10 mile. We offer implications for planners and
- 55 public officials.
- 56

57 Introduction

- 58 Much has been written about the association between rail transit and residential property values
- 59 but less has been written about the association with respect to other property values and nearly
- 60 none about the association with respect to retail property values. We help close this gap. We
- 61 begin with a literature review. This is followed by our research design, study area and data which
- 62 we apply to metropolitan Dallas and Denver. After presenting our results we offer implications
- 63 for planning and public policy.64

65 Literature

- 66 The basic theory of urban economics (1, 2) can be summed up as follows: as a location becomes
- 67 more central to economic activity in a region, demand for such location increases and through
- the market bidding process land becomes more valuable and development becomes more intense.
- 69 Central business district (CBD) location is an obvious example of this. Assuming Von Thünen's
- 70 unfettered plain (3), land values and development intensity tend to fall at a declining rate from the
- 71 CBD. But areas outside CBDs can also enjoy accessibility advantages. This occurs when
- transportation investments confer more efficient accessibility to non-CBD nodes than elsewhere
- outside the CBD. The result should be more intensive development around those. Where those
- 74 investments are highways, congestion often follows thereby undermining efficiencies (4, 5).
- 75 Transit, as an "uncongestible" transportation option (6) can restore accessibility efficiencies,
- 76 leading to yet more intensive development.
- 77
- 78 But do all types of urban land uses react similarly transit station proximity? In an
- 79 important meta-analysis of studies through the middle 2000s, Debrezion, Pels and
- 80 Rietveld (7) identified variations between land uses. Like Bartholomew and Ewing (8),
- 81 they note that most studies of transit-station effects on property values address residential
- 82 property values and most of them address single family values presumably because
- 83 data available for those properties are more plentiful than for other land uses. There were
- 84 about an equal number of studies on attached residential and office properties, but very
- 85 few for other land uses. We will not review the details of their findings except to observe
- that, generally, the literature on industrial, hospitality (principally hotels), and retail
- 87 property value with respect to transit station proximity is small.
- 88
- 89 Generally, most prior studies have assume perhaps as an article of faith that urban land uses
- 90 will cluster mostly within the first one-quarter mile and a few out to about one-half mile.
- 91 Emerging research is challenging the half-mile-circle mantra. For instance, Petheram et al. (9)
- 92 found that apartments capitalized light rail transit station proximity to about 1.25 miles in Salt

distance from the Hiawatha light rail transit stations in Hennepin County, Minnesota to extent 94 95 0.9 mile. For metropolitan Dallas and Denver, however, we (11) find the office rent premium to 96 extend about two miles from light rail transit stations though three-quarters of the premium 97 dissipates at about two-thirds mile. 98 99 We find only one relevant study estimating the rent premium on the association between 100 rail transit proximity and retail properties. Cervero and Duncan (12, 13) find that retail 101 land use value increases substantially within 200 feet of light and commuter rail transit 102 stations, perhaps 167 percent higher than distances beyond 200 feet in San Diego 103 County, California. Our study contributes to knowledge about whether and the extent to 104 which there is an association between retail land uses and, in particular, light rail transit 105 station proximity. We apply our inquiry to metropolitan Dallas and Denver. 106 107 **Research Design, Study Area, Model and Variables** 108 We extend work of others including Ko and Cao by evaluating the retail rent premium associated 109 with light rail transit station proximity in metropolitan Dallas and Denver. We chose those systems for four reasons. First, they are among the oldest LRT systems in the US. The Dallas 110 111 Area Rapid Transit (DART) system began LRT service in 1996 while metropolitan Denver's 112 Regional Transportation District began operating its FasTracks LRT in 1994. Only Portland's 113 (1986), Sacramento's (1987) and San Diego's (1981) LRT systems are older. 114 115 Second, unlike Portland, Sacramento and San Diego, DART and FasTracks serve metropolitan areas that are largely sprawling metropolises undeterred by terrain (the Rocky Mountains are 116 117 tens of miles away from downtown Denver) and policy (neither explicitly contains urban 118 development). 119 120 Third, they are among the nation's largest LRT systems. In 2012, DART had 60 stations and 121 nearly 100,000 daily passengers while FasTracks had 46 stations and nearly 90,000 daily 122 passengers. 123 124 Fourth, their sheer size allow for sufficient data on office rents to undertake hedonic analysis (as 125 we discuss below). Indeed, our study area includes the central counties of Dallas and Denver as 126 well as Arapahoe and Jefferson counties in Colorado. It is thus the largest study area of any study 127 of its kind. 128 129 We employ the following hedonic model in our analysis: 130 131 $\mathbf{R}_i = \mathbf{f}(\mathbf{B}_i, \mathbf{S}_i, \mathbf{C}_i, \mathbf{L}_i)$ 132 133 where: 134 135 R is the market rent per square foot for property *i*; 136 137 B is the set of building attributes of property *i*; 138

Lake County, Utah. Ko and Cao (10) found office and industrial rent premiums with respect to

- 139 S is the set of socioeconomic characteristics of the vicinity of property *i*;
- 140
- 141 C is a composite measure of urban form of the vicinity of property *i*; and
- 142143 L is a set of location attributes of property *i*.
- 144

Our dependent variable, **R** or rent per square foot, and independent variables comprising **B**,
building attributes, come from CoStar, with permission. Through proprietary access during fall

147 2012, we were able to collect an inventory of all retail structures within the study area including

their address, square feet, occupied and vacant space to derive the vacancy rate, stories, effective

149 age (by the later of the construction or renovation year), and weighted average contract rent per

150 square foot though we do not have lease terms for individual tenants. These variables include:

151

152 Socioeconomic data, **S**, come from either the 2010 census (for percent census tract population

- that is not White non-Hispanic) or the 2012 5-year American Community Survey (for census
- 154 tract median household income).
- 155

156 C is a unique variable which measures urban form from most sprawled/diffused/disconnected to 157 most compact/integrated/connected at the level of the census tract. This index places urban 158 sprawl at one end of a continuous scale and compact development at the other. The original 159 index was developed in 2002 for metropolitan areas and counties (14, 15). In a recent study, the 160 compactness indices were refined and updated to 2010 for metropolitan areas, urbanized areas, 161 counties and census tracts and all are posted on a National Institutes of Health website (16).¹ For census tract indices, Ewing and Hamidi used the same methodology and the same type of 162 variables as in larger area analyses. They extracted principal components from multiple 163 164 correlated variables using principal component analysis and transformed the first principal 165 component to an index with the mean of 100 and a standard deviation of 25. Because the number 166 of component variables is greater for street accessibility than land-use mix, and greater for land-167 use mix than development density, the resulting index gives more weight to street accessibility than mix, and to mix than density. This is not unintentional, since the built environment-travel 168 169 literature suggests that density is the least important of the three D variable types (17). Given that 170 retail land uses that depend especially on accessibility this is an appropriate composite variable to include. 171

172

Finally, L, the set of location variables, measures the distance of the centroid of each parcel to
the center of central business district of Dallas or Denver, the nearest entrance onto a limited
access highway and its quadratic term, and distance to the nearest LRT station and its quadratic

- term. Distances are measured in miles.
- 177

178 Although the CoStar retail building database is the most comprehensive available from any

source, only about a quarter of the retail properties include rent. The reason is that most firms

- 180 either own the buildings they use and do not rent space to other tenants, or tenants have long-
- term exclusive tenancy agreements with property owners. Nonetheless, with more than 700 retail
- 182 properties comprising more than 36 million square feet, we believe our analysis will reveal
- 183 central tendencies helping to clarify whether and the extent to which LRT station proximity

¹ http://gis.cancer.gov/tools/urban-sprawl Accessed July 28, 2014.

- 184 confers rent premiums on retail property.
- 185

186 **Results**

187 Table 1 reports results of linear ordinary least squares regression separately for Dallas, Table 2 188 reports results for and Denver, and Table 3 reports combined results. For all models, the 189 coefficients of determination are modest but reasonable given overall performance including 190 expected outcomes with respect to the explanatory variables. The correlation matrices (not 191 reported for brevity) did not reveal problematic correlations, and autocorrelation was not 192 detected. 193 194 **INSERT TABLE 1 ABOUT HERE** 195 **INSERT TABLE 2 ABOUT HERE** 196 **INSERT TABLE 3 ABOUT HERE** 197 198 In all regressions, the building structure variables performed reasonably. The incremental size of 199 a building had no effect on rents suggesting no marginal advantage in larger over small size. The

200 floor area ratio was positive and significant indicating more intensely developed retail sites

201 conveyed higher premiums. It would seem that less parking generates higher rents. On the other

202 hand, the number of stories in a building may depress rent at the margin as floors above (or

below) main levels likely carry goods that sell at lower revenue volumes per square foot than the

204 main level. Increasing vacancy rates reduced mean rents while decreasing effective age increased 205 rents at the margins.

205 206

The socioeconomic variables for Dallas had expected results as increasing median household incomes were associated with increasing while increasing shares of population that were not

209 White Non-Hispanic were associated with decreasing rents. Signs were reversed in Denver

though not significant for income. When both markets are pooled, signs are as expected with

- 211 acceptable levels of statistical significance.
- 212

The Compactness Index was also positive in all regression equations. While this is a composite variable, it suggests that on the whole the market is willing to pay more for locations that are more densely occupied by jobs and people, more integrated in terms of land use mix, and have well-connected streets compared to other locations.

210

The CBD distance location variable performs as expected. In the individual Dallas and Denver regressions, coefficients of the first order and quadratic transformations of the variable measuring distance to the nearest limited access highway entrance had the correct signs though

- in the pooled analysis both were negative but nonetheless consistent with distance-decayexpectations.
- 222 223

224 Of interest to us is the extent to which office rents are affected by proximity to LRT stations and

if so how far away. In the Dallas regression, the coefficients had the expected signs but they

were also just out of range of statistical significance at the 0.10 level of the 1-tailed t-test (since

227 directions of association are predicted). In the Denver equation, the LRT distance-decay

variables also had the expected signs but only the quadratic transformation was significant. In the

229 pooled regression both distance-decay terms had the expected signs and were significant at the

- 230 0.01 level of the 1-tailed t-test. Differentiating the coefficients and then setting for zero we solve
- for the distance threshold. For the pooled markets we estimate the threshold extends about 0.30
- 232 or less than one-third mile.
- 233

234 Implications

- 235 Similar to Cervero and Duncan (12, 13), we find a much tighter distance-decay relationship
- between LRT station proximity and retail rents compared to other land uses. For example, one of
- 237 our recent studies find that apartment land uses capitalize LRT distance up to 1.25 miles away
- (9). We also find office rent premiums to extend in the range of two miles away (11). Ko and
- 239 Cao (10) find a combination of office and industrial rents to extent nearly a mile away. While we
- estimated that half the office premium dissipated after one-half mile and three-quarters dissipated
- after two-thirds mile, in this analysis we find that half the retail premium dissipates at about 0.06 mile while three questions of the mercine dissipates at about 0.10 mile
- 242 mile while three-quarters of the premium dissipates at about 0.10 mile.
- 243
- For decades, planners and public officials have assumed that the largest share of market
- responsiveness to transit stations occurred within the first 0.25 mile and the rest out to about 0.50
- 246 mile. Emerging analysis is relaxing those narrow bands for apartments and office land uses, the
- 247 premiums for which can extend well beyond a mile with half or more of the premium found
- within the first one-half mile. In contrast, this study finds a much tighter distance threshold with
- respect to retail land uses, perhaps only within the first 0.10 mile.
- 250

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- 254 Front Regional Council, the Mountainland Association of Governments, and the University of
- Arizona. We also give special acknowledgement to CoStar for allowing us to evaluate its
- 256 proprietary data on market rents for retail properties in the Dallas and Denver study areas. Our
- 257 views do not necessarily reflect those of our sponsors.
- 258 259

Table 1

- Hedonic Regression Results for Retail Rent Premium with Respect to LRT Station Distance, Dallas

		Std Err	
Variable	Coefficient	of Coef.	t-score p
Constant	-191.242	30.092	-6.355 .01
Gross Leasable Square Feet	0.000	0.000	0.365
Floor Area Ratio	1.945	0.793	2.452 .01
Stories	-1.018	0.598	-1.704 .05
Vacancy Rate	-0.036	0.009	-3.983 .01
Effective Year Built	0.105	0.015	6.825 .01
Median Household Tract Income	0.054	0.012	4.417 .01
Percent Not White Non-Hispanic	-0.032	0.015	-2.076 .05
Compactness Index	3.798	0.615	6.173 .01
Distance from CBD, miles	-0.144	0.077	-1.876 .05
Distance from Interchange, miles	-3.264	1.02	-3.201 .01
Square Distance from Interchange	0.873	0.372	2.347 .01
Distance LRT Station, miles	-1.266	0.687	0.277
Squared Distance LRT Station	1.161	0.138	0.031
R Square	0.289		
Adjusted R Square	0.272		
Std. Error of the Estimate	5.170		
F	17.096		
Sig. F	0.000		
Observations	562		
Degrees of Freedom	548		
Durbin-Watson	1.884		

Table 2

Hedonic Regression Results for Retail Rent Premium with Respect to LRT Station Distance, Denver

		Std Err	
Variable	Coefficient	of Coef.	t-score p
Constant	-242.017	55.035	-4.398 .01
Gross Building Square Feet	0	0	-0.206
Floor Area Ratio	0.427	0.493	0.865
Stories	0.492	1.219	0.404
Vacancy Rate	-0.052	0.016	-3.198 .01
Effective Year Built	0.126	0.028	4.506 .01
Median Household Income	0.009	0.025	0.347
Percent Not White Non-Hispanic	0.236	0.086	2.737 .01
Compactness Index	-1.319	1.176	-1.122
Distance from CBD, miles	0.353	0.202	1.746 .05
Distance from Interchange, miles	0.534	0.508	1.052
Square Distance from Interchange	-0.167	0.069	-2.425 .01
Distance from Nearest LRT Station	-0.608	0.508	-1.199
Squared Distance from Nearest LRT	0.071	0.048	1.471 .05
R Square	0.298		
Adjusted R Square	0.242		
Std. Error of the Estimate	4.671		
F	5.263		
sig. F	0.000		
Observations	175		
Degrees of Freedom	161		
Durbin-Watson	1.750		

Table 3

- Hedonic Regression Results for Retail Rent Premium with Respect to LRT Station Distance, Dallas and Denver

Verieble	Coofficient	Std Err	4
Variable	Coefficient	of Coef.	t-score p
Constant	-184.1570	26.4120	-6.9720 .01
Gross Leasable Square Feet	0.0000	0.0000	-0.0650
Floor Area Ratio	1.3320	0.4350	3.0650 .01
Stories	-0.2860	0.5240	-0.5460
Vacancy Rate	-0.0360	0.0080	-4.6040 .01
Effective Year Built	0.1000	0.0130	7.4600 .01
Median Household Tract Income	0.0550	0.0110	5.0730 .01
Percent Not White Non-Hispanic	-0.0310	0.0140	-2.1100 .05
Compactness Index	2.6520	0.5200	5.0960 .01
Denver	-1.8540	0.7010	-2.6450 .01
Distance from CBD, miles	-0.0690	0.0700	-0.9850
Distance from Interchange, miles	-0.5120	0.3330	-1.5370 .10
Squared Distance from Interchange	-0.1050	0.0580	-1.8240 .01
Distance LRT Station	-0.7930	0.2700	-2.9380 .01
Squared Distance LRT Station	0.1200	0.0320	3.7960 .01
R Square	0.2450		
Adjusted R Square	0.2310		
Std. Error of the Estimate	5.1720		
F	16.1720		
Sig. F	0.0000		
Observations	737		
Degrees of Freedom	722		
Durbin-Watson	1.807		

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