

ARTICLE

Waiting time to radiotherapy as a prognostic factor for glioblastoma patients in a scenario of medical disparities

Tempo de espera para a radioterapia como um fator prognóstico em pacientes com glioblastoma em um cenário de disparidades médicas

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ABSTRACT

Objective: To evaluate the effect of waiting time (WT) to radiotherapy (RT) on overall survival (OS) of glioblastoma (GBM) patients as a reliable prognostic variable in Brazil, a scenario of medical disparities. **Method:** Retrospective study of 115 GBM patients from two different health-care institutions (one public and one private) in Brazil who underwent post-operative RT. **Results:** Median WT to RT was 6 weeks (range, 1.3-17.6). The median OS for WT \leq 6 weeks was 13.5 months (95%CI, 9.1-17.9) and for WT $>$ 6 weeks was 14.2 months (95%CI, 11.2-17.2) (HR 1.165, 95%CI 0.770-1.762; $p = 0.470$). In the multivariate analysis, the variables associated with survival were KPS ($p < 0.001$), extent of resection ($p = 0.009$) and the adjuvant treatment ($p = 0.001$). The KPS interacted with WT to RT (HR 0.128, 95%CI 0.034-0.476; $p = 0.002$), showing that the benefit of KPS on OS depends on the WT to RT. **Conclusion:** No prognostic impact of WT to RT could be detected on the OS. Although there are no data to ensure that delays to RT are tolerable, we may reassure patients that the time-length to initiate treatment does not seem to influence the control of the disease, particularly in face of other prognostic factors.

Keywords: glioblastoma, radiotherapy, waiting time, delay, prognosis, survival.

RESUMO

Objetivo: Avaliar o efeito do tempo de espera (TE) até radioterapia na sobrevida global de pacientes com glioblastoma como um fator prognóstico confiável. **Método:** Estudo retrospectivo de 115 pacientes com glioblastoma, que foram submetidos à radioterapia pós-operatória, em dois serviços diferentes no Brasil (um público e outro privado). **Resultados:** Mediana de TE para radioterapia foi de 6 semanas (variação, 1,3-17,6). A mediana de sobrevida para TE \leq 6 semanas foi de 13,5 meses (IC95%, 9,1-17,9) e para TE $>$ 6 semanas foi de 14,2 meses (IC95%, 11,2-17,2) (HR 1,165, 0,770-1,762; $p = 0,470$). Na análise multivariada, as variáveis associadas à sobrevida foram performance status ($p < 0,001$), extensão da ressecção ($p = 0,009$) e tratamento adjuvante ($p = 0,001$). **Conclusão:** Não se observou impacto prognóstico para TE até a radioterapia na sobrevida. Diante de outros fatores prognósticos, é possível assegurar de que o espaço de tempo até a radioterapia não parece influenciar o controle da doença.

Palavras-chave: glioblastoma, radioterapia, tempo de espera, atraso, prognóstico, sobrevida.

The investigation of the impact of delaying radiotherapy (RT) on the outcomes of various tumors - particularly breast and head-and-neck - has generated no clear results. Nevertheless, there is considerable evidence to support that delay of RT may have an adverse effect on patients' outcomes^{1,2,3}.

Glioblastoma (GBM) patients have been the subject of many clinical investigations aiming to determine the effects of waiting time (WT) for initiation of RT^{4,5,6,7,8,9,10,11,12,13,14,15,16}. Much of the interest in seeking a presumable negative impact of the delay for this kind of tumor is based on its

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median overall survival (OS) that rarely exceeds 14 months¹⁷. So far, available reports are still controversial, although most of the data show no evidence to justify delaying RT beyond 6 weeks¹⁸.

Due to multiple factors, Brazil faces a shortage of RT resources and, according to the most recent data, only 65.9% of the RT demand is met. Nationwide and regardless of tumor type, the mean WT to RT is 113.4 days (almost 4 months) and, barely 15% of the patients initiate RT before 4 weeks following diagnosis^{19,20}. This unfortunate reality and the well-known Brazilian medical disparities have provided the perfect scenario to evaluate a cohort of GBM patients regarding the effect of RT delay on survival. Therefore, this study aimed to evaluate the impact of WT to RT as a reliable prognostic variable in 115 GBM patients from two different health-care institutions in the country.

METHOD

We conducted a retrospective study of adult patients (≥ 18 years old) who had a newly diagnosed GBM according to World Health Organization classification (WHO, 2007)²¹ and were treated with RT after surgery, between January 2003 and December 2011, from two Brazilian hospital databases: *Hospital Israelita Albert Einstein* (HIAE), a private-practice service, and *Hospital São Paulo - Universidade Federal de São Paulo* (HSP-UNIFESP), a public healthcare center.

The medical records of 132 patients were reviewed in order to obtain relevant data: gender, age at diagnosis, Karnofsky Performance Status (KPS), date of first symptom, date of death, date and extent of neurosurgery based on postoperative CT or MRI (within the first 48 hours), and characteristics of RT and chemotherapy. We did not investigate the causes of delay in RT. We excluded 17 patients due to missing data: 3 had no information on the extent of resection, 8 had no information about the adjuvant treatment and, 6 because no starting date of RT was reported. The remaining 115 patients (76 from HSP-UNIFESP and 39 from HIAE) formed our study cohort.

Patients underwent neurosurgical procedure aiming for maximal safe resection whenever possible. All analyzed subjects underwent 3D localized external beam RT using treatment planning systems based on the contrast-enhancing lesion shown on pre-operative contrast-enhanced TC and/or T2/FLAIR sequence MRI. For the purpose of this study, we recorded only the maximum dose actually delivered to the tumor bed at the time of the initial RT, typically 60 Gy in 30 fractions.

At HSP-UNIFESP, before 2009, the patients received Carmustine (BCNU) 200 mg/m² at 6 weeks intervals starting 6 weeks after RT. Since 2009, Temozolomide (TMZ) became available and patients could be treated according to

the EORTC-NCIC protocol²². At HIAE, all patients were treated according to EORTC-NCIC protocol²². The patients who underwent concomitant and adjuvant TMZ were categorized as “RT concurrent with chemotherapy”; those who received BCNU were defined as “RT and sequencing chemotherapy”.

The ethics review board of both institutions approved this study.

Statistical analysis

Data were described using absolute and relative frequencies for categorical data. Quantitative data were described using median and range, due to skewness. OS was calculated from time of diagnosis until death or last follow up (cut-off date October 17, 2012).

The waiting time between first symptoms and neurosurgery (WT to NS) was calculated using the estimated date of first symptoms to the day of the neurosurgical intervention (biopsy or surgical resection) and was categorized based on its median. The WT to RT was calculated in weeks between the date of the first neurosurgery and the starting date of RT. We categorized WT to RT based on its median into two groups: RT delay ≤ 6 weeks (WT ≤ 6) and RT delay > 6 weeks (WT > 6). Data from the groups were compared by using Pearson's Chi-square or Fisher's test for categorical data and Mann-Whitney *U*-test for quantitative variables. A *p*-value of < 0.05 was considered significant.

Survival curves were constructed according to the Kaplan-Meier method and compared between groups using log-rank test to explore relationships between well-recognized prognostic factors (age, KPS, extent of resection, adjuvant treatment) and survival in the univariate analysis. Data from two diverse institutions were evaluated and we observed a significant association between WT to RT and institution. In order to avoid multicollinearity, only the WT to RT was explored as a prognostic factor.

A conditional stepwise proportional hazard analysis (Cox-regression model) was used to identify independent predictors of survival. The variables that achieved a *p*-value < 0.1 in the univariate analysis were included in the multivariate model. The WT to RT variable was retained in the model because it is the object of the study, although not significant in the univariate analysis. In addition, we explored the interactions among WT to RT and other variables, and the multivariate model was adjusted to include only significant interactions.

The statistical analysis was performed using the statistical softwares R (R Core Team (2012). R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. ISBN 3-900051-07-0, URL <http://www.R-project.org/>) and SPSS (SPSS Inc. Released 2008. SPSS Statistics for Windows, Version 17.0. Chicago: SPSS Inc.).

RESULTS

For the entire cohort, the median age was 57 years (range, 18-83). There were 44 females and 71 males (ratio 1:1.6). The median intervals were: WT to NS 6.7 weeks (range, 0.14-1435) and WT to RT 6 weeks (range, 1.3-17.6). Median final RT dose was 60 Gy (range, 10-66). Median follow-up time was 12.9 months (range, 1.1-58) and estimated OS was 14.1 months (95%CI, 11.6-16.6).

The median interval for WT ≤ 6 group was 4.1 weeks (range, 1.3-6) and 8.6 weeks (range, 6.1-17.6) for the WT > 6 group. Table 1 shows patients' demographics, clinical and treatment characteristics as a function of the WT to RT interval.

We did not observe a deleterious effect of longer WT to RT on survival, as shown in Figure 1. The median OS for WT ≤ 6 was 13.5 months (95%CI, 9.1-17.9) and for WT > 6 was 14.2 months (95%CI, 11.2-17.2) (HR 1.165, 95%CI 0.770-1.762; p = 0.470). We observed a significantly better OS for age < 50 years, KPS ≥ 70%, final total RT dose ≥ 60 Gy and adjuvant treatment including any chemotherapy (Table 2).

We explored the impact of clinical characteristics and adjuvant treatment on OS using a Cox regression model (Table 3). The only significant clinical variable associated with survival was KPS (p < 0.001). Regarding treatment characteristics, the extent of resection (p = 0.009) and the adjuvant treatment (p = 0.001) had also an association with OS.

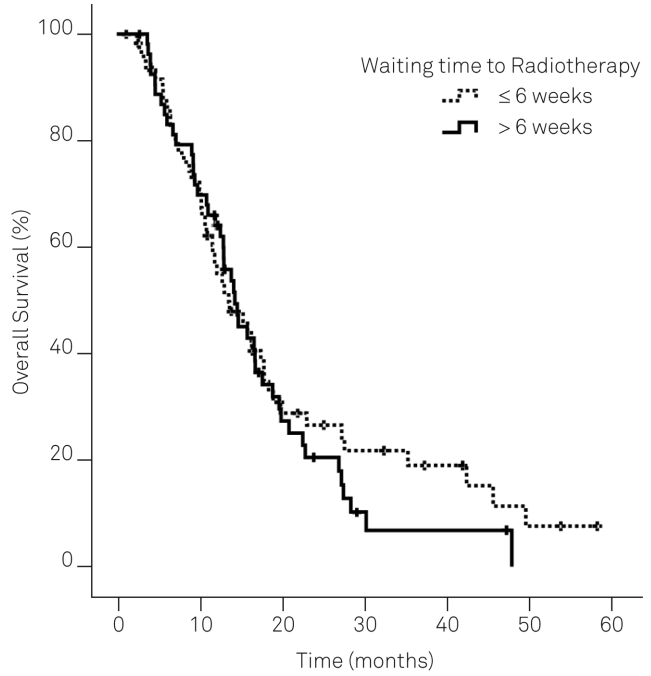


Figure 1. Overall survival according to waiting time from neurosurgery to radiotherapy for glioblastoma patients (p = 0.470).

We also evaluated the potential interactions between patient's characteristics and WT to RT. The only variable that interacted with WT to RT was KPS. Figure 2 shows that the effect of KPS on OS depends on the WT to RT.

Table 1. Demographics, clinical and treatment characteristics of patients according to the WT to RT interval.

	WT ≤ 6 61 (53%)	WT > 6 54 (47%)	p
Gender			0.159
Male	34 (55.7%)	37 (68.5%)	
Age (years)			0.071
< 50	14 (23.0%)	21 (38.9%)	
≥ 50	47 (77.0%)	33 (61.1%)	
KPS (%)			0.423
< 70	11 (23.4%)	9 (17.0%)	
≥ 70	36 (76.6%)	44 (83.0%)	
Neurosurgery (extent of resection)			0.115
Biopsy	8 (13.1%)	3 (5.6%)	
Partial	20 (32.8%)	27 (50.0%)	
Gross-total	33 (54.1%)	24 (44.4%)	
WT to NS			0.021
≤ 6.7 weeks	33 (61.1%)	21 (38.9%)	
> 6.7 weeks	21 (38.9%)	33 (61.1%)	
Final total RT dose (Gy)			0.050
< 60	15 (26.3%)	23 (44.2%)	
≥ 60	42 (73.7%)	29 (55.8%)	
Adjuvant treatment			0.004
RT Only	13 (21.3%)	15 (27.8%)	
RT and sequencing chemotherapy	11 (18.0%)	22 (40.7%)	
RT concurrent with chemotherapy	37 (60.7%)	17 (31.5%)	

WT: Waiting time; RT: Radiotherapy; KPS: Karnofsky performance status; WT to NS: Waiting time between first symptoms and neurosurgery.

Table 2. Overall survival according to patients' characteristics.

	Median survival time in months (95%CI)	HR	(95%CI)	p
Age (years)				0.032
< 50	18.8 (15.0-22.6)			
≥ 50	12.5 (10.1-14.9)	1.637	(1.042-2.571)	
KPS (%)				< 0.001
< 70	9.0 (5.1-12.9)			
≥ 70	15.8 (13.3-18.2)	0.377	(0.220-0.648)	
Neurosurgery (extent of resection)				0.096
Biopsy	11.5 (9.7-13.3)			
Partial	15.7 (11.3-20.1)	0.613	(0.297-1.264)	
Gross-total	14.1 (11.8-16.4)	0.464	(0.224-0.962)	
WT to NS				0.295
≤ 6.7 weeks	15.7 (12.7-18.8)			
> 6.7 weeks	12.9 (10.0-15.7)	1.252	(0.822-1.908)	
Final total RT dose (Gy)				0.042
< 60	11.0 (8.2-13.7)			
≥ 60	16.2 (13.5-19.0)	0.641	(0.418-0.985)	
Adjuvant Treatment				< 0.001
RT Only	8.1 (5.2-11.0)			
RT and sequencing chemotherapy	15.8 (13.8-17.7)	0.282	(0.163-0.487)	
RT concurrent with chemotherapy	17.7 (15.2-20.2)	0.231	(0.136-0.394)	

WT: Waiting time; RT: Radiotherapy; KPS: Karnofsky performance status; WT to NS: Waiting time between first symptoms and neurosurgery; HR: Hazard ratio; CI: confidence interval.

DISCUSSION

In this study of 115 GBM Brazilian patients, delaying RT longer than 6 weeks did not affect the OS and may not be credited as a reliable prognostic factor (HR 1.323, 95%CI 0.731-2.393; p = 0.355). The problem of delaying RT in GBM treatment has been debated more extensively because of the fast-growing nature of this tumor¹³ and its ability to

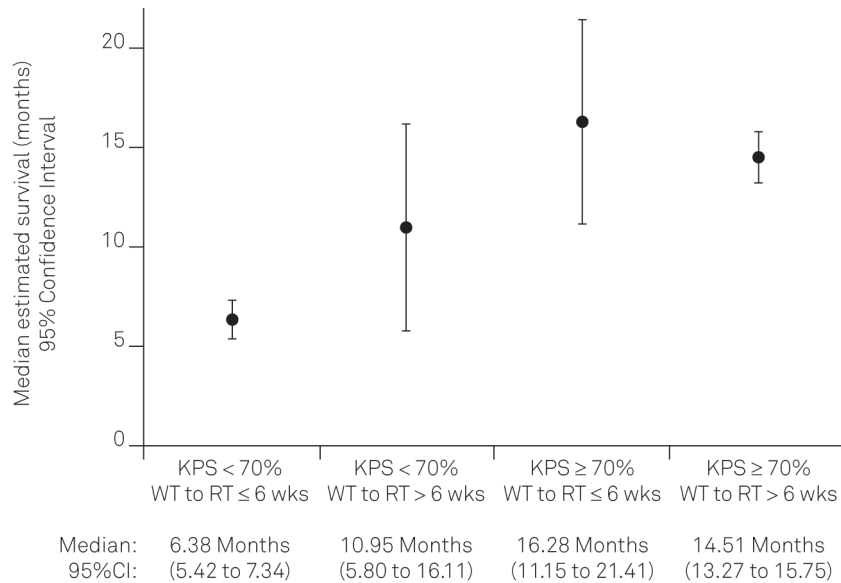
invade surrounding tissues, making tumor control much more difficult over time^{23,24}. To date, data published on this theme are considered controversial, partially because prior studies presented diverse methodologies and patients with different characteristics^{4,5,6,7,8,9,10,11,12,13,16,18}.

In line with our data, previous reports failed in detecting any disadvantage in delaying RT. In 2000, Do et al.⁴ in the first published attempt to address the issue, reported

Table 3. Multivariate analysis.

	HR	(95%CI)	p
WT to RT			0.355
≤ 6 weeks	1		
> 6 weeks	1.323	(0.731-2.393)	
Age (years)			0.837
< 50	1		
≥ 50	0.939	(0.518-1.705)	
KPS (%)			< 0.001
≤ 70	1		
< 70	6.794	(2.799-16.49)	
Neurosurgery (extent of resection)*			
Gross-total	1		
Partial	1.174	(0.712-1.937)	0.530
Biopsy	4.382	(1.703-11.278)	0.002
Final total dose of RT (Gy)			
≥ 60	1		
< 60	1.176	(0.708-1.954)	0.531
Adjuvant treatment†			
RT concurrent with chemotherapy	1		
RT and sequencing chemotherapy	1.184	(0.651-2.154)	0.580
RT Only	3.259	(1.713-6.202)	< 0.001

*Global p-value = 0.009; †Global p-value = 0.001. WT: Waiting time; RT: Radiotherapy; KPS: Karnofsky performance status; HR: Hazard ratio; CI: Confidence interval.



WT: Waiting time; RT: Radiotherapy; KPS: Karnofsky performance status.

Figure 2. Median estimated survival time according to WT to RT and KPS.

no detrimental impact on survival when longer intervals between neurosurgery and RT occurred (HR 1.00, 95%CI 0.99-1.02; $p = 0.09$). The main limitation of that study was the inclusion in the same cohort of grade 3 and grade 4 gliomas, which are currently known to present different behaviors. Other authors^{6,8,9,10,12}, already working in the era of TMZ concomitant with RT, support the result of no significant effect of longer WT to RT on OS in GBM patients. Likewise, using a huge cohort extracted from patients enrolled in the Radiation Therapy Oncology Group (RTOG) protocols, and accounting for close to 3,000 GBM patients, Blumenthal et al.⁷ found no reduction in survival associated with RT delay, within a temporal limit of 6 weeks. Remarkably, median survival was significantly greater for those with the longest WT to RT (> 4 weeks) than the group presenting a shorter WT to RT (≤ 2 weeks) ($p < 0.001$). Some authors^{14,15,16} reported comparable unexpected results when patients submitted to a longer WT to RT had also a longer survival. Similarly, we observed the same tendency for better survival for WT > 6 (median 14.2 months), although not significant ($p = 0.470$). These patients (WT > 6) had no remarkable advantage regarding age, KPS and, treatment characteristics if compared to the WT ≤ 6 group. No clear radiobiological mechanisms can be proposed to explain such phenomenon, although physician selection bias of poor performance patients has already been pointed out as an alternative reason⁷. Some authors corroborate this hypothesis suggesting that better performing patients wait longer to RT^{5,9,16,18}. In our cohort, the only clinical variable that interacted with WT to RT was KPS (HR 0.128, 95%CI 0.034-0.476; $p = 0.002$), showing that the benefit of KPS on OS depends on

the WT to RT. In fact, as depicted in Figure 2, when patients with KPS $< 70\%$ were submitted to RT in ≤ 6 weeks they showed the worst OS (6.4 months; range, 5.4-7.3). This is a particular potential bias for some retrospective studies published so far, since poor prognosis patients seem to be first in line for treatment.

Three studies reached conclusions opposing our data. Irwin et al.⁵ were the first to report a negative impact of the delay in RT for GBM patients. The study reported a mean WT to RT of 5.01 weeks and, each additional week of delay increases the risk of death by 8.9% (HR 1.089, 95% CI 1.020-1.161; $p = 0.010$). The major limitations of the study were the inclusion of two pathology subtypes (grade 3 and 4 gliomas) and the great variability of RT dose used. Only 30.8% of the patients received a final dose of ≥ 60 Gy and some patients had their RT dose chosen based upon the performance status. In addition, in their multivariate analysis, the RT dose was found to positively influence survival. Our patients were more homogeneous since the final total RT dose between the groups was similar ($p = 0.05$), and all subjects underwent the same RT technique over the period, targeting a final dose of 60 Gy. As for Valdivieco et al.¹¹ study, 107 GBM patients submitted to a complete surgical resection were analyzed and showed WT to RT ≤ 6 weeks as an independent survival predictor (HR 0.42, 95%CI 0.06-0.96; $p = 0.043$). However, some limitations must be pointed out: first, the WT to RT cut-off of 6 weeks was chosen based on historical data and not on its median, causing an imbalance between the groups; second, the restricted inclusion criterion of completed resected patients may limit the generalizability of their results; third, significantly more patients (86%,

$p = 0.0014$) in the RT delay ≤ 6 weeks subgroup received some kind of chemotherapy, creating a potential bias since the use of chemotherapy (TMZ or nitrosureas) was associated with longer survival in the multivariate analysis. In our study, we decided to use the TMZ concomitant with RT as the reference value in the multivariate analysis, aiming to minimize the interpretation bias and we did not detect any advantage when comparing TMZ to the use of sequential chemotherapy (BCNU). In addition, the use of adjuvant chemotherapy (BCNU or TMZ) was equivalently allocated between our subgroups (WT ≤ 6 weeks 78% x WT > 6 weeks 71%). Lastly, in an impressive cohort of 345 TMZ treated GBM patients, Spratt et al.¹⁶ demonstrated a detrimental effect of delaying RT greater than 6 weeks (HR 3.76, 95%CI 1.01-14.57; $p = 0.05$) for a subset of patients with known O6-methylguanine DNA methyltransferase (MGMT) status; although only 45.8% of their cohort had MGMT status. No detrimental impact on OS was found (HR 1.68, 95%CI 0.94-3.06; $p = 0.08$) when the entire cohort was analyzed, including unknown MGMT status patients.

Although we could not retrieve the causes of delaying RT, there is an interest in understanding what would justify the WTs to RT since it is well known that the delay may cause distress and sometimes impact the quality of life due to persisting symptoms²⁴. Many authors argue that the inevitable reason for increasing WTs is the failure to reconcile RT to its demand, creating a continuous imbalance^{1,2,3}. Even developed countries face this problem^{25,26,27,28,29}, which may be even more profound in primarily government funded health care systems. In Brazil, despite recent government efforts¹⁹, the scenario is still worrisome and the growing demand for RT has created an inefficient system where only 65.9% of the demand is met and long WTs are commonplace with a median WT to RT of 113.4 days (almost 4 months)^{19,20}, regardless of tumor type. Indeed, our data reveals a median interval for WT > 6 group of 8.6 weeks, but in this subset, WT as long as 17.6 weeks could be found. Another relevant aspect of the situation is the lengthy admission process into the health care system, which may be reflected in the WT to NS (median of 6.7 weeks). Most patients (61.1%) who waited longer for neurosurgery (> 6.7 weeks) also had a longer WT to RT (> 6 weeks) ($p = 0.021$). This may be seen as an indication of failure to obtain adequate care when early symptoms first occur. We are aware though that the date of the first symptom may be a source of error, which may add bias to these results.

Our study has some limitations. First, its retrospective nature may subject the study to common drawbacks such

as patient uncontrolled selection bias and missing data. Thus, our results may not apply in other temporal and spatial contexts. To minimize the selection bias we controlled our analysis considering well-known prognostic factors. It should be noted that a randomized controlled trial to finally elucidate this important debate might not be possible due to ethical issues^{1,2,3,18}. Second, we included patients from two diverse health-care institutions and the WT to RT was significantly related to the institution. However, we must emphasize that to avoid multicollinearity, only the WT to RT was explored as a prognostic factor. Additionally, we reported before that the medical assistance scenario was not an independent predictor of survival³⁰ and, by electing patients from two diverse healthcare settings, we aimed to reflect the Brazilian reality. Third, we included only patients submitted to RT at some point of their treatment. Lastly, the long time period of the study may have affected the final results due to the diverse adjuvant treatment approaches. To minimize this potential bias, as stated earlier, we decided to use the RT concurrent with chemotherapy as the reference treatment in the multivariate analysis and observed that no advantage was found when comparing this regimen to the use of sequencing chemotherapy (BCNU) (HR 1.184, 95%CI 0.651-2.154; $p = 0.580$). Therefore, it is unlikely that the addition of TMZ in the treatment arsenal caused any bias to the outcome of our sample.

Our results echo much of the previous published data and no prognostic impact of WT to RT could be detected on the survival of GBM patients. Although there are no data to ensure that delays to RT are tolerable, we may reassure patients that the control of the disease does not seem to be influenced by the time-length to initiate treatment, particularly when we consider the power of other prognostic factors. In any case, we should not lose sight of the increasing imbalance between the RT demand and supply and the imperative need of a carefully planned strategy by the healthcare system to handle this challenge.

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