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Citation for published version:

Größbacher, V, Buková, K, Lawrence, AB, Špinka, M & Winckler, C 2020, 'Discriminating spontaneous locomotor play of dairy calves using accelerometers', *Journal of Dairy Science*, pp. 1866-1873. https://doi.org/10.3168/jds.2019-17005

Digital Object Identifier (DOI):

10.3168/jds.2019-17005

Link: Link to publication record in Edinburgh Research Explorer

Document Version: Peer reviewed version

Published In: Journal of Dairy Science

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Running head: DISCRIMINATING CALF PLAY USING ACCELEROMETERS

3	Discriminating spontaneous locomotor play of dairy calves using accelerometers
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ABSTRACT

Play behavior is a promising welfare indicator in dairy calves as it decreases in negative 27 situations such as pain or hunger and increases in positive contexts such as in appropriate social 28 environment. Directly measuring play is time consuming as it is performed in irregular bouts 29 and can be inconsistent over days. To facilitate automatic recording of play, previous studies 30 31 fitted tri-axial accelerometers to the hind legs of calves, measuring the velocity of movements in large arenas, and reported high correlations between vertical axis peak duration and the 32 duration of locomotor play. The current study aimed at validating accelerometers for recording 33 spontaneous locomotor play in the calves' home-pens over longer periods of time. Data were 34 collected from 48 Holstein Friesian calves at either four or eight weeks of age, housed in groups 35 of three in pens of 10 m². Acceleration at the vertical axis of the hind leg was recorded at a rate 36 of 1 Hz. One active time period for each calf was randomly selected (mean duration \pm SD = 34 37 \pm 9 min). From video of the corresponding time period, frequency of locomotor play events 38 39 consisting of run, turn and buck/buck-kick was recorded using behavior sampling. Combined counts of play events were highly correlated ($r_p = 0.91$) with counts of peaks in acceleration. 40 However, for calves with higher levels of locomotor play, this method underestimated the 41 extent of play. Alternatively, run, turn and buck events obtained from video were transformed 42 into a binary response by creating intervals of 10s and then classifying each 10s interval as 43 44 comprising events of play (PLAY) or not comprising events of play (NOPLAY). The corresponding accelerometer data for all 10s periods, equaling 10 consecutive readings each, 45 were classified into PLAY or NOPLAY with quadratic discriminant analysis. 79% of periods 46 47 with locomotor play were correctly classified. Counts of observed play intervals correlated with the counts of play periods from accelerometers at $r_p = 0.87$, but the discriminant analysis 48 consistently overestimated play. In conclusion, accelerometer measurements at 1 Hz (in 1 s 49 50 intervals) and at the vertical axis alone cannot be used to exactly quantify absolute levels of locomotor play in the home-pen. However, counts of peak accelerations can provide a rough 51

- 52 estimate of inter-individual differences in play events and discriminant analysis can be used as
- a proxy for one-zero sampling of inter-individual differences in locomotor play.
- 54 Key words: automated measuring, acceleration, behavior classification, dairy calf

INTRODUCTION

56 In the past decade, accelerometers have found increasing application in farm animal research. The use of accelerometers facilitates data collection as automatic recording can circumvent 57 time and resource intensive behavior observations. In calves, accelerometers have been 58 primarily employed to record general activities. For example lying, standing or locomotion of 59 dairy calves have been recorded using accelerometers to detect early behavioral signs of 60 61 respiratory diseases (Swartz et al., 2017) and of neonatal diarrhea (Sutherland et al., 2018b). Moreover acceleration measurements have been used to quantify lying and standing when 62 studying effects of social housing on weaning (Overvest et al., 2018) and determining effects 63 64 of different disbudding methods on lying behavior (Sutherland et al., 2018a). While the 65 accuracy of recording general activities from accelerometers is high, the validation of recording specific behaviors such as feeding and ruminating is still in progress (e.g. Roland et 66 67 al., 2018). In calves, play behavior is a promising indicator to assess both compromised welfare, e.g. 68

pain after disbudding (Mintline et al., 2013) or hunger after weaning off milk reduced play 69 (Krachun et al., 2010, Miguel-Pacheco et al., 2015), and enhanced welfare, e.g. group housing 70 71 increased play (Valníčková et al., 2015). However, calves perform spontaneous play for only 72 a few minutes per day (Jensen et al., 2015) at irregular intervals (Fraser and Duncan, 1998). Thereby quantification of play from observation is usually accomplished either through 73 continuous recording of durations or events (e.g. Jensen et al., 2015, Miguel-Pacheco et al., 74 75 2015) or through one-zero sampling of the presence of play in certain sample intervals (e.g. Valníčková et al., 2015). These challenges associated with measuring play behavior raise the 76 77 interest in automatic recording techniques. In previous studies, accelerometers were used to automatically record locomotor play of calves, however the recordings were conducted in 78 large arenas and for a short time only. Rushen and de Passillé (2012) found correlations of up 79 80 to $r_s = 0.88$ between the duration of running and the sum of total acceleration in all three axes

81	and Luu et al. (2013) found correlations of up to $r_p = 0.98$ between the duration of locomotor
82	play (running plus jumping/kicking) and the sum of the percent of peaks (3 g or higher) of all
83	axes. In both studies acceleration was recorded at a high rate of 33 Hz and in all three axes,
84	limiting the recording duration to 10 min due to the memory capacity of the accelerometers
85	(HOBO Pendant G Acceleration Data Logger, Onset Computer Corporation, Pocasset, MA,
86	USA). With the intention of assessing a longer recording duration, Luu et al. (2013) simulated
87	a sampling frequency of 1 Hz by taking every 33rd acceleration reading on one axis only and
88	found a correlation of $r_p = 0.92$ between the percent of peaks on the vertical axis and the
89	duration of locomotor play. In order to test the practical application of measuring spontaneous
90	locomotor play over the duration of many hours, the current study aimed to validate the use of
91	accelerometers to measure play behavior in the home-pens of calves at low recording rates.
92	Specifically, our objectives were using recordings at 1 Hz on the vertical axis (1) to test
93	whether counts of peak accelerations can accurately estimate events of locomotor play and (2)
94	to determine whether classifying periods of acceleration readings into PLAY/NOPLAY can
95	reliably measure play behavior recorded by one-zero sampling from video.

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MATERIAL AND METHODS

All data were collected at the Netluky Research Station at the Institute of Animal Science in
Prague/Czech Republic between August 2016 and April 2017. The study was approved by the
Institutional Animal Care and Use Committee of the Institute of Animal Science in Prague
and the Czech Central Committee for Protection of Animals, Ministry of Agriculture (permit
number 27356/2016-MZE-17214).

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104 Animals and Housing

The 48 Holstein Friesian-calves (20 female, 28 male) reported on here were a subset of a
larger study using 72 calves. They were housed in an uninsulated barn with wind-shields in 24

groups of three. Pens were 10.1 m^2 with a straw-bedded lying area of 4.2 x 1.4 m and a 107 108 concrete activity and feeding area of 3.5 x 1.2 m. Calves entered group-housing at an average age of 13.3 ± 3.1 days (mean \pm SD) with groups entering the experiment consecutively. Calf 109 allocation to groups was balanced for sex, age and weight. For the purpose of another 110 experiment calves were fed either 6 liters of milk daily throughout the experiment or they 111 received 9 liters per day at week 4 and the provision continuously increased to 12 liters at 112 113 week 6. All calves received three milk meals per day in teat buckets. All calves received 3 liters of milk in the morning. Calves with a low milk allowance received 1.5 liters of milk per 114 meal at midday and in the evening. Calves with a high milk allowance received 3 liters 115 116 continually increasing to 4.5 liters of milk per meal at midday and in the evening. Calves had 117 ad libitum access to water, concentrates and hay offered in buckets. Among the 72 calves, two focal calves per group were randomly selected with observations of one calf taking place at 4 118 weeks and the other at 8 weeks of age. Calves weighed 57.5 \pm 5.7 kg (mean \pm SD) at 4 weeks 119 and 88.3 ± 12.4 kg at 8 weeks. 120

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122 Acceleration measurements

Accelerometers (HOBO Pendant G Acceleration Data Logger, Onset Computer Corporation,
Pocasset, MA, USA; product specifications are described in detail in Luu et al. (2013)) were
fitted to the rear side of both hind legs of calves using elastic cohesive bandages.
Accelerometers were attached vertically to the leg such that the x-axis was perpendicular to
the ground. The accelerometers were set to measure readings on the vertical axis at a rate of 1
Hz (1 sample/s), allowing recording of acceleration every second for 18.1 hours. Acceleration
was recorded from 05.00 until 23.04 on two consecutive days per testing week with the

accelerometer on the right leg recording day 1 and on the left leg recording day 2. Calves

131 wore accelerometers for approximately 66 hours per testing week (from the evening before

132 until the morning after the testing days). Programming of accelerometers was performed using

an optical infrared base station with USB interface and the HOBOware Pro Software (Version
3.7.8; Onset Computer Corporation, Pocasset, MA, USA) with the starting time set in
advance.

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137 Behavior observation

Behavior of calves was video recorded for 48 hours per testing week using one camera per 138 pen (VCC-HD2300P, Sanyo, Japan; FW2220R-Z, Dahua Technology Co., China; HDC-139 SD99, Panasonic, Japan) and infrared radiators (RM50-AI-50, Raytec, UK; LIR-T80 and 140 LIR-T60, IR LAB Surveillance Tech, Taiwan). Based on the graphic display of downloaded 141 142 acceleration values, using the plot-function of the HOBOware Pro Software, lying and active 143 phases could be clearly distinguished. Therewith one activity bout of approx. 30 min was selected for each calf. Activity bouts were selected in a time span between 05.00 and 20.00 144 when accelerometer recordings were available and video recordings allowed easy distinction 145 of behaviors due to daylight hours. The week and day of the selected activity bout was 146 randomized for each focal calf. Selection of activity bouts was balanced across different times 147 of the day and start time of selected bouts ranged from 06.04 until 19.23. The duration of 148 149 selected activity bouts was 34.3 ± 9.2 min (mean \pm SD). For individual recognition calves 150 were marked across their backs and sides with animal marking sticks. Behaviors categorized as locomotor play are described in Table 1. The criterion interval for halts in between running 151 events was set to 1 second based on visual assessment of a log survivorship plot. Events of 152 153 locomotor play behavior were continuously recorded by one person using the Mangold INTERACT video analysis software (Version 16.1.5.8). Intra-observer-reliability was 154 measured from 3 randomly selected activity bouts of 41.7 ± 8.0 min each (mean \pm SD) 155 assessed two times. A Wexler's ratio was calculated from the number of agreements (i.e. the 156 number of locomotor play events that were correctly scored within one second in both 157 158 recording sessions x 2) divided by the number of possible agreements (i.e. the total number of

locomotor play events scored in both sessions; e.g. used in Wathan et al. 2015). Wexler's ratio
was assessed for each activity bout individually, with an average agreement ratio of 0.84.
Continuous recording was transformed into one-zero sampling by creating sample intervals of
and classifying them according to presence or absence of locomotor play events within
the interval.

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165 Data analysis

All statistical analyses were performed in SAS 9.4. We analyzed the acceleration dataaccording to two methodologies:

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Peak acceleration method (PEAK). We used Pearson correlations to assess the strength of 169 association between counts of peak measurements of acceleration and counts of observed 170 171 locomotor play events. Counts of peak accelerations were calculated for different upper and lower thresholds of acceleration values in steps of 0.1 g (e.g. counts of values \geq 3.2 g, 3.1 g, 172 3.0 g,... and $\leq -3.2 g, -3.1 g, -3.0 g,...$). Pearson correlations of all 1056 combinations (32) 173 thresholds of ≥ 0 g times 31 thresholds of < 0 g) of counts of peaks applying different upper 174 175 and lower thresholds with counts of locomotor play were calculated. Therewith, the best 176 combination of threshold values of acceleration to predict locomotor play was identified as counts of peaks of \geq + 1.6 g and \leq - 3.0 g. While the distribution of data was right-skewed and 177 did not visually conform to the assumption of normality for parametric measures of 178 179 association, the data distribution was unimodal. Three outliers were visually detected using a Cook's Distance plot, though no outlier with leverage was visually identified in the outlier 180 and leverage diagnostics (leverage and studentized residuals). 181

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Classifier method (CLASS). We used quadratic discriminant analysis to predict a categorical
response (Kuhlenkasper and Handl, 2017), i.e. the occurrence of locomotor play in each

period (10 s fragment of observations) based on classifiers (predictor variables describing 185 186 acceleration values in each period (James et al., 2015)). As discriminant analysis requires two sets of data, one set to train the discriminant function and one set to test its predictions, we 187 divided the recorded activity bouts in half. Therefore the accelerometer data were combined 188 to 10 s periods, resulting in 10 measurements per period. Subsequently the periods were 189 alternately allocated to a training data set or a testing data set (testing data set: n=48, mean 190 191 number of periods \pm SD = 102.3 \pm 26.9). The presence (PLAY) or absence (NOPLAY) of locomotor play in each period was identified from video observation and used as the gold 192 standard. For each period the following metrics were calculated as classifiers derived from the 193 194 original value (OV) or change in values ($CV = x_i - x_{i-1}$) e.g. minimum, maximum, mean, median, quartiles, variance, total sum; a full list is provided in Supplemental Table S1. 195 Relevant classifiers were then visually preselected from boxplots of PLAY and NOPLAY 196 197 from the training data set when the interquartile range of NOPLAY was low with little to no overlap with PLAY and when outliers were not widely dispersed. A quadratic discriminant 198 function was then developed with classification probabilities based on the proportional 199 occurrences of how often PLAY and NOPLAY were scored in the training data set, i.e. 97% 200 201 of periods displaying NOPLAY and 3% of periods displaying PLAY. With the testing data set 202 the predictive abilities of the discriminant function were assessed. Discriminant functions with different combinations of classifiers were tested and the combination of classifiers with 203 the highest sensitivity and specificity was selected. The relevant classifiers included in the 204 205 final discriminant function are displayed in Table 2. Discriminant analysis assumes a Gaussian distribution from observations of each class (James et al., 2015). The present data of 206 207 the values of classifiers could not be assumed to be normally distributed nor could data be transformed to fit the underlying assumptions of normality. We were able to circumvent this 208 issue by dividing the data set into two halves, a training data set and a testing data set. 209

Therewith the performance of the discriminant function was not contingent on the datadistribution and could be independently verified.

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Comparison of PEAK and CLASS. In order to directly contrast the outcome of the two 213 methodologies on the basis of the same set of data, we calculated PEAK and CLASS with the 214 215 testing data set only (n = 48; mean duration \pm SD = 17.1 \pm 4.6 min). To assess the strength of 216 association between the measures of acceleration and the observed locomotor play, a Pearson correlation of counts of peaks resulting from the PEAK method and counts of observed 217 locomotor play events was calculated. Likewise, a Pearson correlation of counts of PLAY 218 219 periods resulting from the CLASS method with counts of observed locomotor play intervals 220 from one-zero sampling was calculated. In order to assess the magnitude of disagreement and facilitate the detection of trends, we produced Bland-Altman plots. The plots depict the 221 222 average of the acceleration measure and the observation on the x-axis and the difference 223 between the acceleration measure and the observation on the y-axis (Altman and Bland, 1983). Bland-Altman plots were produced for both methodologies of analysis and compared 224 visually. 225

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RESULTS

228	When assessing play by continuous recording of frequencies, calves performed 5.3 events of
229	locomotor play per 30 min observation period (SD = 7.3 ; range = $0 - 27$). The Pearson
230	correlation with counts of peaks of \geq + 3.0 g and \leq - 3.0 g from the corresponding
231	accelerometer data, as described by Luu et al. (2013), was 0.83 (P<0.01). However, we
232	attained the highest correlation with counts of locomotor play when using counts of peaks of
233	\geq + 1.6 g and \leq - 3.0 g (r _p = 0.91, P > 0.01; Figure 1). The respective scatter plot (Figure 1)
234	illustrates a strong linear relationship of both measurements, but an unequal rate of increase of
235	counts of peaks with counts of play is noticeable. The Bland-Altman plot (Figure 2) further

emphasizes the uneven distribution across the range of locomotor play as higher counts of 236 237 locomotor play events were increasingly underestimated by the peak acceleration method, demonstrating that the number of play events and the number of accelerometer peaks did not 238 directly correspond to each other, i.e., they are not on the same scale. The mean deviation of 239 240 peak measurements from observed play events amounted to - 1.90 ± 4.42 . Alternatively, when recording locomotor play with one-zero sampling, calves performed play 241 242 in 2.7 periods per observation (SD = 3.5; range = 0 - 16). From the accelerometer data, we estimated the number of play periods using the classifier method with the outcome displayed 243 as contingency table (Table 3). It follows that CLASS overestimates the number of PLAY 244 245 periods. CLASS achieved a precision of 0.95 (= proportion of correctly classified periods), a sensitivity of 0.79 (= proportion of correctly classified true positives) and a specificity of 0.96 246 (= proportion of correctly classified true negatives). Counts of PLAY periods identified with 247 248 CLASS highly correlated with counts of observed PLAY periods recorded from video ($r_p =$ 0.87; P < 0.01; Figure 3). The scatter plot (Figure 3) illustrates a strong linear relationship of 249 both measurements but indicates an intercept and concomitant overestimation of PLAY 250 periods by the CLASS method. The number of accelerometer-identified PLAY periods 251 252 surpasses the number recorded visually by 3.65 ± 2.42 periods; nonetheless the Bland-Altman 253 plot (Figure 4) shows an evenly distributed deviation of the two measurements across the range of counts of PLAY periods. 254

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DISCUSSION

With this study we aimed at providing an approach to automatically record locomotor play of calves in their home-pen and for long durations using acceleration measurements. In previous studies accelerometers have been validly used to record durations of lying and standing in calves (Bonk et al., 2013, Swartz et al., 2016). Similarly in the current study we were able to easily distinguish between lying and standing on the vertical axis, with values of lying

fluctuating around 0 g and values of standing around - 1 g, depending on the position of the 262 263 hind leg. Therefore with - 1 g as the center of fluctuation, measuring play with peaks of $\geq +$ 1.6 g and \leq - 3.0 g is sensible. We reason that peaks had not reached + 3.0 g, as reported by 264 Luu et al. (2013), because the smaller dimensions of the home-pens in comparison with a 265 266 large arena did not permit calves to consistently reach accelerations of a similarly high level. Thus small spaces can restrict the magnitude of movement and also fragment the occurrence 267 268 of play (Jensen et al. 1998). Nevertheless we cannot draw conclusions on any space allowance between our home-pen and the arena of Luu et al. (2013) as this was not part of our 269 investigation. Moreover locomotor play consists of rapid motions of the hind legs for short 270 271 durations and is often nested within short time intervals. Therefore recordings at 1 Hz and on 272 one axis may be too infrequent to accurately capture locomotor play events in the home-pen, resulting in the unequal increase and accretive underestimation of higher frequencies of play 273 274 events of the PEAK-method, as visualized in the Scatter plot and Bland-Altman plot. Nevertheless, the high correlation of peak accelerations and observed play events elucidates a 275 strong link between the two recording methods. Thus, while the PEAK-method cannot record 276 the duration of play in absolute terms, it can produce an approximate estimation of play levels 277 278 and allows the comparison of relative differences between calves in standard housing 279 conditions.

280 In the CLASS method, we used the accelerometer data to simulate the one-zero observational method by merging the recordings to 10 s periods, thus ensuring the use of repeated measures 281 282 and circumventing the need to count individual peaks above/below a certain threshold. This allowed us to view acceleration values in context, integrated with the values preceding and 283 following them. We derived classifiers from combined values e.g. mean of two highest values 284 or variance to mathematically describe the 10 acceleration measures per period and highlight 285 the differences between PLAY and NOPLAY. Thereby we classified brief time spans 286 287 according to the presence or absence of locomotor play within these 10 seconds. The use of

original individual acceleration values e.g. the mere minimum or maximum value would have 288 289 resulted in a lower sensitivity to correctly identify PLAY periods. Such an approach has been previously successfully implemented in accelerometer validation regarding sheep gait, 290 describing periods with relative frequencies of integers e.g. the number of high frequency 291 acceleration readings between - 4 and - 3 per period divided by all readings of the period 292 (Radeski and Ilieski, 2017). Other studies described periods using movement metrics e.g. 293 294 mean, variance and inverse coefficient of variation (Watanabe et al., 2008) or signal 295 magnitude area, average intensity and average entropy (Barwick et al., 2018). However, in these studies acceleration was recorded at a higher rate. Recording at a higher rate would have 296 297 also allowed for classifying shorter periods. For example Radeski and Ilieski (2017) recorded 298 at 33 Hz and classified periods of three seconds.

CLASS correctly discriminated 79% of periods with an occurrence of locomotor play, but at 299 300 the same time overestimated play by approximately 200% (out of 304 play periods identified by CLASS, 102 periods were true positives and 202 periods were false positives). Similarly 301 other accelerometer models consistently overestimated locomotor behavior e.g. Swartz et al. 302 (2016) overestimated stepping by 18% and Trénel et al. (2009) consistently overestimated 303 304 moving activity with a ratio of probability of correct negatives to correct positives of 7.57 305 (PV - = 0.98, PV + = 0.13). Thus while the number of classified play periods is strongly 306 associated with counts of observed play intervals, overall the classifier method overestimates locomotor play in absolute terms and produces an intercept by adding 3.7 play periods to each 307 308 observation. However, the Bland-Altmann plot shows a rather consistent and evenly distributed deviation across the range of number of play intervals observed without indication 309 310 of a directed effect. Hence, while CLASS cannot accurately measure locomotor play, it can be used as a proxy. After factoring in the consistent overestimation, it assesses the number of 311 play periods close to the scale of one-zero sampling and thus allows comparing absolute 312 313 differences between individuals. With these results we offer a feasible approach to assess

spontaneous locomotor play in home-pens of calves using an affordable and commercially 314 315 available accelerometer model for durations of many hours or perhaps even days. Nevertheless these results may only be valid for the housing conditions investigated and 316 further studies are needed to validate this approach under e.g. different space allowances. A 317 318 prerequisite to classify periods with discriminant analysis is to use shorter subsets of behavior recordings as a training data set. In the current study only active periods of animals were 319 320 included in the analysis. In order to apply the classifier method to the full data set it is necessary to either preselect only active periods or to include lying bouts in the training data 321 set. Therewith it is feasible to train the discriminant function with the selected classifiers and 322 323 thereafter to apply it to the entire recordings of acceleration. 324 We must stress that our proposed approach with recording at a frequency of 1 Hz can only be used as an approximate estimation of locomotor play. A higher level of accuracy could be 325 326 achieved by increasing the rate of recording. Measuring acceleration at the highest rate (33 Hz) allowed de Passillé et al. (2010) to measure the interstep interval and accurately 327 distinguish between different gait patterns. Radeski and Ilieski (2017) were able to achieve 328 high accuracy in classifying 3 s periods of walking, trotting and galloping in sheep with 329 330 discriminant analysis, when recorded at a rate of 33 Hz. In the current study the recording rate 331 was limited by its data storage capacity, however Le Roux et al. (2018) achieved a 469-fold reduction in memory requirement when classifying lying, standing and walking on the 332 accelerometer rather than storing raw data. Thus, the proposed approach is easily applicable 333 334 and inexpensive with the available resources, however there are numerous options to improve the accuracy of recording by availing technical advancements. 335

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CONCLUSION

Using the peak acceleration method, the acceleration of calves' hind legs measured at a rate of1 Hz can be used to obtain an approximate estimation of inter-individual differences in the

340	occurrence of locomotor play events. Quadratic discriminant analysis can replace
341	observational one-zero sampling, when based on indirect movement metrics obtained from
342	10-second-periods of raw accelerometer data. This alternative method may be more accurate
343	in quantifying the inter-individual differences in locomotor play of dairy calves in their home-
344	pens as it reveals less biased estimates across different levels of play. If the accurate
345	measurement of absolute levels of behavior is the ultimate aim of automatic recording, a
346	sensor with higher memory capacity must be found.
347	
348	ACKNOWLEDGEMENTS
349	We are grateful to the farm staff at the Netluky Research Station for the care of calves and
350	cooperation with the research team. This study was funded by the Austrian Academy of
351	Sciences and by the grant MZE-RO0718 from the Czech Ministry of Agriculture.
352	
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Table 1. Ethogram of locomotor play derived from Jensen et al. (1998) and Jensen and Kyhn (2000)

Type of locomotor play	Description			
Running	Rapid gait with phase of suspension in the air. Minimum of 2 consecutive suspension movements in a forward direction. Running is counted as a new event after 1 second break.			
Turning	The two forelegs are lifted from the ground and stretched forward, as the forepart of the body is elevated and turned to one side. Movement upwards and sideward for a minimum of 90 degrees. Occurrence is scored during running bouts.			
Bucking/Buck-kicking	Simultaneous lifting of hind legs, claws are raised to a level as high as, or higher than tarsal joints in a standing position. One or both hind legs may be kicked in a posterior or lateral direction. Occurrence is scored during running bouts.			

Table 2. Descriptions and equations of the classifiers included in the final discriminant
 435

function. Means \pm standard deviation of the classifiers are shown for the periods of the testing 436

data set identified as PLAY or NOPLAY from video. OV = original values, CV = change in 437

values, PLAY = period with presence of locomotor play, NOPLAY = period with absence of 438 439 locomotor play

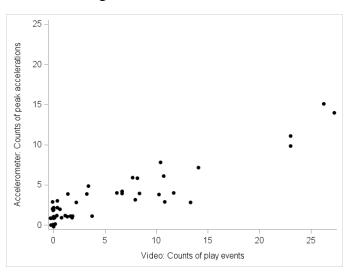
Classifier	Equation	NOPLAY	PLAY
OV: Mean of two highest acceleration measurements	$\frac{\max(x) + \max 2(x)}{2}$	-0.89 ± 0.21	0.04 ± 0.99
OV: Mean of two lowest acceleration measurements	$\frac{\min(x) + \min 2(x)}{2}$	-1.06 ± 0.20	-1.85 ± 0.70
OV: Variance	$\frac{\sum (x-\mu)^2}{10}$	0.03 ± 0.10	0.74 ± 0.80
CV: Maximum of absolute value of change in acceleration measurements	$\max(\Delta x)$	0.26 ± 0.48	2.41 ± 1.45
CV: Mean of change in acceleration measurements	$\frac{1}{10}\sum\Delta x_i$	-0.00 ± 0.03	-0.01 ± 0.12
CV: Total sum of absolute values of change in acceleration measurements	$\sum \Delta x_i $	0.66 ± 1.15	6.96 ± 4.78

- Table 3. Contingency table with number of periods identified with the classifier method
- (CLASS) as PLAY (event of locomotor play occurring in this period) and NOPLAY (no event of locomotor play occurring in this period)

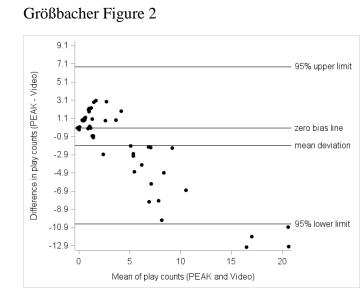
Predicted behavior			
(CLASS)			
NO			
PLAY	PLAY	Sum	
4591	202	4793	
27	102	129	
4618	304	4922	
	(CLASS NO	CLASS NO PLAY PLAY 4591 202 27 102	

- **Figure 1.** Relationship between counts of peak accelerations ($\geq 1.6 g$ and $\leq 3.0 g$; PEAK) and
- 446 counts of locomotor play events observed from video (n = 48 calves). Jitter function was used
- in the graph to make multiple identical values more visible



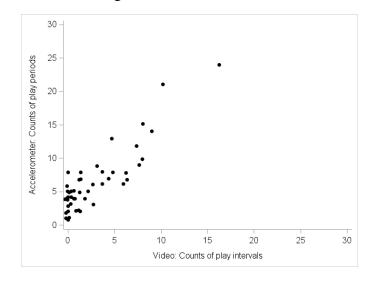


- 450 Figure 2. Bland-Altman Plot of the difference in the assessment of locomotor play recorded
- 451 with accelerometers and video observation compared with the mean of both assessments
- 452 (PEAK = Peak acceleration method; n = 48 calves). Confidence intervals were estimated at
- 453 6.8 at the 95% upper limit and 10.6 at the 95% lower limit. Jitter function was used in the
- 454 graph to make multiple identical values more visible

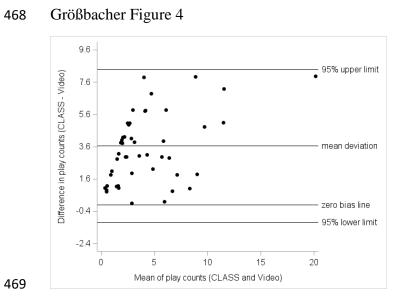


- 457 Figure 3. Relationship between counts of periods with locomotor play identified with
- 458 accelerometers (CLASS = classifier method) and counts of sample intervals with locomotor
- 459 play observed from video (n = 48 calves). Jitter function was used in the graph to make
- 460 multiple identical values more visible

461 Größbacher Figure 3



- 463 Figure 4. Bland-Altman Plot of the difference in the assessment of locomotor play periods
- identified with accelerometers and locomotor play intervals from video observation compared
- with the mean of both assessments (CLASS = Classifier method; n = 48 calves). Confidence
- 466 intervals were estimated at 8.4 at the 95% upper limit and 1.1 at the 95% lower limit. Jitter
- 467 function was used in the graph to make multiple identical values more visible





470 **Supplemental Table S1.** Descriptions and equations of all potential classifiers not included in the final discriminant function. Means ± standard

deviation of the classifiers are shown for the periods of the testing data set identified as PLAY or NOPLAY from video. OV = original values, CV =472 change in values, PLAY = period with presence of locomotor play, NOPLAY = period with absence of locomotor play. Potential classifiers that

473 were preselected and tested but not included in the final discriminant function are marked with 'Yes'

Classifier	Equation	NOPLAY	PLAY	Preselection
OV: Highest acceleration measurement	$\max(x)$	-0.84 ± 0.33	0.44 ± 1.30	Yes
OV: Second highest acceleration measurement	$\max^2(x)$	-0.94 ± 0.12	-0.37 ± 0.84	Yes
OV: Mean of two highest acceleration measurements	$\frac{\max(x) + \max 2(x)}{2}$	-0.89 ± 0.21	0.04 ± 0.99	Yes
OV: Third quartile of acceleration measurements	$x_{Q0.75}$	$\textbf{-0.96} \pm 0.08$	-0.64 ± 0.50	No
OV: Mean of acceleration measurements	$\frac{1}{10}\sum x_i$	-0.98 ± 0.07	-0.93 ± 0.29	No
OV: First quartile of acceleration measurements	$x_{Q0.25}$	-1.00 ± 0.07	-1.24 ± 0.41	No
OV: Mean of two lowest acceleration measurements	$\frac{\min(\mathbf{x}) + \min 2(\mathbf{x})}{2}$	-1.06 ± 0.20	-1.85 ± 0.70	Yes
OV: Second lowest acceleration measurement	min2(x)	-1.02 ± 0.10	-1.52 ± 0.69	Yes
OV: Lowest acceleration measurement	$\min(x)$	-1.11 ± 0.33	-2.19 ± 0.86	Yes
OV: Variance	$\frac{\sum (x-\mu)^2}{10}$	0.03 ± 0.10	0.74 ± 0.80	Yes
OV: Total sum of absolute values of acceleration measurements	$\sum x_i $	9.84 ± 0.63	11.32 ± 2.20	Yes

CV: Highest change in acceleration measurements	$\max(\Delta x)$	0.22 ± 0.43	2.01 ± 1.34	No
CV: Second highest change in acceleration measurements	$\max 2(\Delta x)$	0.07 ± 0.14	0.86 ± 0.81	No
CV: Third quartile of change in acceleration measurements	$\Delta x_{Q0.75}$	0.04 ± 0.07	0.50 ± 0.55	No
CV: Mean of change in acceleration measurements	$\frac{1}{10}\sum\Delta x_i$	-0.00 ± 0.03	-0.01 ± 0.12	Yes
CV: First quartile of change in acceleration measurements	$\Delta x_{Q0.25}$	$\textbf{-0.04} \pm 0.08$	$\textbf{-0.49} \pm 0.53$	No
CV: Second lowest change in acceleration measurement	min2(Δx)	$\textbf{-0.07} \pm 0.15$	$\textbf{-0.88} \pm 0.82$	Yes
CV: Minimum of change in acceleration measurement	$\min(\Delta x)$	$\textbf{-0.23} \pm 0.44$	-2.07 ± 1.33	Yes
CV: Variance of change in acceleration measurements	$\frac{\sum (\Delta x - \mu)^2}{10}$	0.05 ± 0.21	1.56 ± 1.75	Yes
CV: Maximum of absolute value of change in acceleration measurements	$\max(\Delta x)$	0.26 ± 0.48	2.41 ± 1.45	Yes
CV: Total sum of absolute values of change in acceleration measurements	$\sum \Delta x_i $	0.66 ± 1.15	6.96 ± 4.78	Yes