

**THE ANALYSIS OF ROAD TRAFFIC ACCIDENT DATA IN THE
IMPLEMENTATION OF ROAD SAFETY REMEDIAL
PROGRAMMES.**

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DECLARATION

I the undersigned hereby declare that the work contained in this thesis is my own original work and has not previously in its entirety or in part been submitted at any university for a degree.

SIGNATURE

DATE

ABSTRACT

A road safety remedial programme has as an objective the improvement of road transportation safety by applying road safety engineering remedial measures to hazardous road network elements in a manner that will be economically *efficient*.

Since accident data is the primary manifestation of poor safety levels it must be analysed in manner that will support the overall objective of economic efficiency.

Three steps in the process of implementing a road safety remedial programme, that rely on the systematic analysis of accident data, are the identification of hazardous locations, the ranking of hazardous locations and the evaluation of remedial measure effectiveness.

The efficiency of a road safety remedial programme can be enhanced by using appropriate methodologies to measure safety, identify and rank hazardous locations and to determine the effectiveness of road safety remedial measures. There are a number of methodologies available to perform these tasks, although some perform much better than other. Methodologies based on the Empirical Bayesian approach generally provide better results than the Conventional methods. Bayesian methodologies are not often used in South Africa. To do so would require the additional training of students and engineering professionals as well as more research by tertiary and other research institutions.

The efficiency of a road safety remedial programme can be compromised by using poor quality accident data. In South Africa the quality of accident data is generally poor and should more attention be given to the proper management and control of accident data.

This thesis will report on, investigate and evaluate Bayesian and Conventional accident data analysis methodologies.

ABSTRAK

Die doel van 'n padveiligheidsverbeteringsprogram is om op die mees koste effektiewe manier die veiligheid van onveilige padnetwerkelemente te verbeter deur die toepassing van ingenieursmaatreëls.

Aangesien padveiligheid direk verband hou met verkeersongelukke vereis die koste effektiewe implementering van 'n padveiligheidsverbeteringsprogram die doelgerigte en korrekte ontleding van ongeluksdata.

Om 'n padveiligheidsverbeteringsprogram te implementeer word die ontleding van ongeluksdata verlang vir die identifisering en priortisering van gevaarkolle, sowel as om die effektiwiteit van verbeteringsmaatreëls te bepaal.

Die koste effektiwiteit van 'n padveiligheidsverbeteringsprogram kan verbeter word deur die regte metodes te kies om padveiligheid te meet, gevaarkolle te identifiseer en te prioritiseer en om die effektiwiteit van verbeteringsmaatreëls te bepaal. Daar is verskeie metodes om hierdie ontledings te doen, alhoewel sommige van die metodes beter is as ander. Die 'Bayesian' metodes lewer oor die algemeen beter resultate as die gewone konvensionele metodes. 'Bayesian' metodes word nie in Suid Afrika toegepas nie. Om dit te doen sal addisionele opleiding van studente en ingenieurs vereis, sowel as addisionele navorsing deur universiteite en ander navorsing instansies.

Die gebruik van swak kwaliteit ongeluksdata kan die integriteit van 'n padveiligheidsverbeteringsprogram benadeel. Die kwaliteit van ongeluksdata in Suid Afrika is oor die algemeen swak en behoort meer aandag gegee te word aan die bestuur en kontrole van ongeluksdata.

Die doel van hierdie tesis is om verslag te doen oor 'Bayesian' en konvensionele metodes wat gebruik kan word om ongeluksdata te ontleed, dit te ondersoek en te evalueer.

There is no such thing as an accident.
What we call by that name is the effect
Of some cause which we do not see.

VOLTAIRE

Carriages without horses shall go,
And accidents fill the world with woe.

Prophecy attributed to Mother Shipton (17th century)

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TABLE OF CONTENTS

	Page
• ABSTRACT	iii
• ABSTRAK	iv
• ACKNOWLEDGEMENTS	vi
• TABLE OF CONTENTS	vii

CHAPTER 1

1. INTRODUCTION

1.1	THE ROAD SAFETY SITUATION IN SOUTH AFRICA	1-1
1.2	ACCIDENT CAUSATION	1-1
1.3	ROAD SAFETY MANAGEMENT	1-3
1.4	THE ROLE OF ROAD SAFETY ENGINEERING	1-4
1.5	ROAD SAFETY REMEDIAL PROGRAMMES	1-6
1.6	THE ROLE OF ACCIDENT DATA ANALYSIS	1-10
1.7	STUDY OBJECTIVES	1-11

CHAPTER 2

2. INFORMATION REQUIREMENTS

2.1	INTRODUCTION	2-1
2.2	ACCIDENT DATA	2-2
2.2.1	Data quality	2-2
2.2.2	Reportability	2-2
2.2.3	Accident classification	2-4
2.2.4	Accident reporting	2-5
2.3	ROAD NETWORK INFORMATION	2-14
2.4	TRAFFIC VOLUME INFORMATION	2-15
2.4.1	Accuracy of traffic counts	2-16
2.5	INFORMATION MANAGEMENT	2-21
2.5.1	Database design	2-22
2.5.2	Data extraction and reduction	2-25
2.6	SUMMARY and CONCLUSION	2-32

CHAPTER 3

3. MEASURING SAFETY

3.1	INTRODUCTION	3-1
3.2	DEFINING SAFETY	3-2
3.3	MEASUREMENT THEORY	3-3
3.3.1	Validity and reliability	3-4
3.3.2	Measurement error	3-4
3.4	ACCIDENT MEASURE	3-6
3.5	EXPOSURE	3-9
3.5.1	Road segments	3-11
3.5.2	Intersections	3-13
3.6	AN ALTERNATIVE APPROACH TO SAFETY ESTIMATION	3-14
3.7	SAFETY PERFORMANCE FUNCTIONS	3-15
3.7.1	Road segments	3-16
3.7.2	Intersections	3-22
3.7.3	The argument averaging problem	3-26
3.7.4	The function averaging problem	3-32
3.8	SUMMARY and CONCLUSION	3-34

CHAPTER 4

4. SAFETY MEASUREMENT METHODOLOGIES

4.1	INTRODUCTION	4-1
4.2	HISTORICAL RATE METHOD	4-2
4.3	GENERIC CLASS METHOD	4-6
4.4	COMBINING THE GENERIC AND HISTORICAL RATE METHODS	4-7
4.5	THE EMPIRICAL BAYESIAN APPROACH	4-8
4.5.1	Introduction	4-8
4.5.2	Theoretical framework	4-9
4.5.3	Estimating $E(m)$ and $VAR(m)$	4-10
4.5.4	Applying Bayes theorem	4-15
4.5.5	Performance	4-21
4.5.6	The multivariate regression method without a reference group	4-24
4.5.7	The multivariate regression method – a more coherent approach	4-27
4.5.8	Statistical inference	4-31
4.6	SUMMARY and CONCLUSIONS	4-37

CHAPTER 5

5. IDENTIFICATION AND RANKING OF HAZARDOUS LOCATIONS

5.1	INTRODUCTION	5-1
5.2	PROGRAMME DESIGN	5-2
5.3	COMPONENTS OF IDENTIFICATION METHODS	5-4
5.3.1	Single sites	5-4
5.3.2	Sections/Segments/Routes	5-6
5.4	IDENTIFICATION OF HAZARDOUS LOCATIONS	5-8
5.4.1	Conventional identification methods	5-9
5.4.2	Bayesian identification methods	5-25
5.4.3	Comparison of identification methods	5-29
5.5	RANKING METHODOLOGIES	5-33
5.5.1	Accident rate/number ranking method	5-33
5.5.2	Bayesian safety estimate method	5-35
5.5.3	Potential Accident Reduction index (PAR) method	5-35
5.5.4	Degree of deviation	5-40
5.5.5	Severity method	5-41
5.6	PERFORMANCE OF IDENTIFICATION AND RANKING METHODS	5-44
5.7	SUMMARY and CONCLUSION	5-46

CHAPTER 6

6. THE EVALUATION OF ROAD SAFETY REMEDIAL MEASURES

6.1	INTRODUCTION	6-1
6.2	CONCEPTUAL FRAMEWORK	6-2
6.2.1	Predicting accidents	6-3
6.2.2	Statistical framework	6-3
6.2.3	Methodological framework	6-4
6.3	SIMPLE BEFORE AND AFTER METHODOLOGY	6-5
6.3.1	Statistical analysis	6-9
6.3.2	Study design	6-11
6.3.3	Performance of before-and-after methodology	6-17
6.4	ACCOUNTING FOR THE EXPOSURE EFFECT	6-19
6.4.1	Statistical analysis	6-19
6.5	BEFORE-AND-AFTER WITH COMPARISON GROUP METHOD	6-22
6.5.1	Statistical framework	6-23
6.5.2	Study design	6-29
6.6	THE EMPIRICAL BAYES APPROACH	6-35
6.6.1	The simple before-and-after procedure	6-35
6.6.2	Before-and-after with comparison group method	6-39

6.6.3	The multivariate regression ‘time-series’ approach	6-40
6.6.4	Allowing for changes in exposure	6-44
6.7	THE TREATMENT EFFECT	6-53
6.7.1	Theoretical framework	6-50
6.8	SUMMARY and CONCLUSION	6-52

CHAPTER 7

7. MULTIVARIATE REGRESSION MODELLING

7.1	INTRODUCTION	7-1
7.2	MULTIPLE REGRESSION ANALYSIS – AN INTRODUCTION	7-2
7.3	ORDINARY MULTIPLE LINEAR REGRESSION	7-5
7.3.1	Model form	7-5
7.3.2	Assumptions	7-7
7.3.3	Model estimation	7-7
7.3.4	Performance	7-8
7.4	GENERALISED LINEAR MODELLING	7-9
7.4.1	Theoretical framework	7-9
7.4.2	Poisson model formulation	7-10
7.4.3	Log-linear model estimation	7-15
7.4.4	Estimation techniques	7-17
7.4.5	Estimation tools	7-18
7.4.6	Model evaluation	7-18

7.5	SUMMARY and CONCLUSION	7-21
-----	------------------------	------

CHAPTER 8

8.	SUMMARY AND CONCLUSIONS	8-1
----	-------------------------	-----

CHAPTER 9

9.	RECOMMENDATIONS	9-1
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LIST OF REFERENCES

APPENDICES

APPENDIX A1 : Experimental data

APPENDIX A2 : Comparing the ability of the conventional and Bayesian methods to accurately measure the true level of safety.

APPENDIX A3 : Efficiency assessment of identification methods.

APPENDIX A4 : Assessment of evaluation methods.

APPENDIX B1: The method of sample moments.

APPENDIX B2 : Identification of hazardous locations.

APPENDIX B3 : Ranking of hazardous locations

LIST OF TABLES

- Table 1.1 :** 1998 South African accident statistics.
- Table 1.2 :** Unit cost of accidents by severity and status.
- Table 2.1 :** Coefficient-of-variation for different stratification methods and counting periods.
- Table 2.2 :** Coefficient-of variation values for *mother/daughter* and *direct estimation* methods.
- Table 2.3 :** Coefficient-of-variation values per traffic category and counting period.
- Table 2.4 :** Components of database design.
- Table 3.1 :** Regression parameters of Safety Performance Functions (Persaud and Musci , 1995).
- Table 3.2 :** Regression parameters of Safety Performance Functions (Persaud , 1993).
- Table 3.3 :** Regression parameters of SPF's for signalised intersections (Hauer et al. , 1988).
- Table 3.4 :** Safety Performance Functions (Mountain an Fawaz, 1996).
- Table 3.5 :** Example 3.5 – Expected accident numbers.
- Table 4.1 :** Example 4.1 – Calculations.
- Table 4.2 :** Example 4.2 – Year 1 data – estimation of mean and variance.
- Table 4.3 :** Example 4.3 – Groups of entities – accident number method.
- Table 4.4 :** Example 4.7 – Summary statistics : Conventional and Bayesian methods.
- Table 4.5 :** Example 4.8 – Calculations.
- Table 4.6 :** Example 4.9 – Data and calculations.
- Table 4.7 :** Using the GAMMADIST function of Excel®.
- Table 4.8 :** Incomplete Gamma function values ($\alpha=1$) : 95 % Degree of Confidence.

- Table 4.9 :** Incomplete Gamma function values ($\alpha=1$) : 5 % Degree of Confidence.
- Table 5.1 :** Proposed influence areas of intersection at different speeds.
- Table 5.2 :** Accident number method : Efficiency assessment of different threshold values.
- Table 5.3 :** Probability factor (k) values.
- Table 5.4 :** Cost of accidents in 1988.
- Table 5.5 :** Example 5.3 – Accident data.
- Table 5.6 :** Example 5.4 – Mean, standard deviation and threshold values.
- Table 5.7 :** Example 5.4 – List of hazardous locations from Methods CN1, CR1 and NR.
- Table 5.8 :** Example 5.5 - Method CN2 – threshold values.
- Table 5.9 :** Example 5.5 - Method CN2 – efficiency assessment.
- Table 5.10 :** Example 5.7 – Method CP1 – critical values.
- Table 5.11 :** Example 5.7 - Method CP1 – efficiency assessment.
- Table 5.12 :** Example 5.8 – Method CR2 : Top 10 segments where $R > R_{cr}$.
- Table 5.13 :** Example 5.9 – Method B1 – efficiency assessment.
- Table 5.14 :** Example 5.10 – Method B2 – top 10 hazardous locations.
- Table 5.15 :** Example 5.12 – Comparison of CR1, CR2, B1 and B2 identification methods.
- Table 5.16 :** Example 5.13 - Top 10 sites : Accident Number method.
- Table 5.17 :** Example 5.13 – Top 10 sites : Accident Rate method.
- Table 5.18 :** Example 5.14 – Top 10 sites : Bayesian estimate method.
- Table 5.19 :** Example 5.15 – Top 10 sites according to PAR index.
- Table 5.20 :** Data for comparison of identification and ranking methods (Persaud et al. 1999b).
- Table 6.1 :** The 4-step process.
- Table 6.2 :** Example 6.1 – Annual average accident frequency data.
- Table 6.3 :** The 4-step process for the simple before-and-after procedure.
- Table 6.4 :** Example 6.2 – Accident data.
- Table 6.5 :** Example 6.2 – The 4-step procedure : Calculations.

- Table 6.6 :** Example 6.3 – Before-and-after accident data and calculations.
- Table 6.7 :** Example 6.3 – The 4-step procedure – calculations.
- Table 6.8 :** Minimum expected accidents in before period : 95 % degree of confidence.
- Table 6.9 :** Minimum expected accidents in before period : 99 % degree of confidence.
- Table 6.10 :** Example 6.5 - Annual accident data : Hermanus : 1991 to 1998.
- Table 6.11 :** Example 6.5 - The 4-step procedure – calculations.
- Table 6.12 :** Example 6.5 – The 4-step procedure – calculations.
- Table 6.13 :** Case Study 6.1 – Case study ‘before’ and ‘after’ accident data.
- Table 6.14 :** Case Study 6.1 – The 4-step procedure – calculations.
- Table 6.15 :** Example 6.6 – Results of assessment of conventional identification and evaluation methods.
- Table 6.16 :** The 4-step procedure : Accounting for the exposure effect.
- Table 6.17 :** Example 6.7 – Data and calculations.
- Table 6.18 :** Example 6.7 – The 4-step procedure – calculations.
- Table 6.19 :** Example 6.7 – The 4-step procedure – calculations.
- Table 6.20 :** The 4-step procedure : Comparison group method.
- Table 6.21 :** Example 6.8 – Treatment and comparison group data.
- Table 6.22 :** Example 6.8 - The 4-step procedure – calculations.
- Table 6.23 :** The 4-step procedure – entities with own comparison ratio.
- Table 6.24 :** Example 6.9 – Data and calculations.
- Table 6.25 :** Example 6.9 – The 4-step procedure – calculations.
- Table 6.26 :** Example 6.10 – Data and calculations.
- Table 6.27 :** Example 6.11 – Data and calculations.
- Table 6.28 :** Example 6.11 – Data and calculations.
- Table 6.29 :** Example 6.11 – The 4-step procedure – calculations.
- Table 6.30 :** Comparison group : Minimum sample size : $r_d = 1$ and $\text{var}(\omega)=0$.
- Table 6.31 :** Comparison group : Min. sample size : $r_d = 0.5$ and $\text{var}(\omega)=0$.
- Table 6.32 :** Comparison group : Minimum sample size : $r_d = 2$ and $\text{var}(\omega)=0$.
- Table 6.33 :** Case Study 6.1 – Treatment and comparison group data.
- Table 6.34 :** Case Study 6.1 – The 4-step procedure – calculations.
- Table 6.35 :** The 4-step procedure : Bayesian before-and-after method.

Table 6.36 : Case Study 6.1 – Before-and-after Bayesian estimates

Table 6.37 : Example 6.13 – Assessment of simple before-and-after Bayesian methodology.

Table 6.38 : The 4-step procedure : Bayesian before-and-after with comparison group method.

Table 6.39 : Example 6.14 – Data and calculations.

Table 6.40 : Example 6.14 – Data and calculations.

Table 6.41 : Example 6.14 – The 4-step procedure – calculations.

Table 6.42 : Case Study 6.2 – ‘Before’ and ‘after’ data.

Table 6.43 : Case Study 6.2 – Results of study.

Table 6.44 : Case Study 6.2 – The 4-step procedure – calculations.

Table 6.45 : The extended 4-step procedure – estimating treatment effects.

Table 6.46 : Case Study 6.1 – Data and calculations.

Table 6.47 : Case Study 6.1 – The extended 4-step procedure.

Table 7.1 : Case Study 7.1 – Traffic flows for accident pattern 6.

Table 7.2 : Case Study 7.2 – Results of multivariate regression analysis.

Table A1.1 : Appendix A1 – Summary statistics

Table A1.2 : Appendix A1 – Listing of data for 1000 experimental sites.

Table A2.1 : Appendix A2 – Conventional and Bayesian estimates and deviations from the ‘true’ level of safety (Sites 1 to 50).

Table A2.2 : Appendix A2 - Cumulative frequencies (%) : Conventional, Generic and Bayesian estimates.

Table A3.1 : Appendix A3 – Methods CP1 and CN2 – Identification of hazardous locations. (Sites 1 to 50)

Table A3.2 : Appendix A2 – Method B1 – Identification of hazardous locations. (Sites 1 to 50).

Table A4.1 : Appendix A4 – Hazardous locations – Method CP1 : Period = 3 years.

Table A4.2 : Appendix A4 – Hazardous locations – Method B1 : Period = 3 years

Table B1.1 : Appendix B1 – Western Cape, Class 1, 2-lane rural road network and traffic count information – Method of sample moments.

Table B2.1 : Appendix B2 – Identification of hazardous locations.

Table B3.1 : Appendix B3 - Comparison of ranking procedures.

LIST OF FIGURES

Figure 1.1 : Factors contributing to road traffic accidents.

Figure 2.1 : Accuracy vs. degree of accident reporting.

Figure 2.2 : A database table.

Figure 2.3 : A database table with redundant information.

Figure 2.4 : A database table.

Figure 2.5 : A database table.

Figure 2.6 : Normalised database tables.

Figure 3.1 : Typical categorisation of intersection accidents.

Figure 3.2 : Non-linear Safety Performance Function.

Figure 3.3 : SPF's for single and multi-vehicle accidents on Michigan freeways.

Figure 3.4 : Accident number and rate curves : Exponential SPF.

Figure 3.5 : Single vehicle and multi-vehicle accident SPF's.

Figure 3.6 : Night-time and day-time accident SPF's.

Figure 3.7 : Accident number and rate curves : Quadratic SPF.

Figure 3.8 : Accident no. and rate curves : Hoerl's function with $k = 1$ and 2 .

Figure 3.9 : Intersection accident patterns (Hauer et al. ; 1988)

Figure 3.10 : The argument averaging problem.

Figure 3.11 : Correction factor (w) – Exponential SPF.

Figure 3.12 : Correction factor (w) – Quadratic SPF.

Figure 3.13 : Correction factor (w) : Hoerl's function : $k=1$.

Figure 3.14 : Correction factor (w) : Hoerl's function : $k=2$

Figure 3.15 : The Function Averaging problem.

Figure 4.1 : Example 4.1 - Accuracy of different study periods.

Figure 4.2 : Example 4.1 – Mean deviations for different study periods.

Figure 4.3 : Accident rates at various levels of exposure.

Figure 4.4 : Comparison of *generic* and *5-year historical* methods.

Figure 4.5 : Example 4.7 - Bayesian estimation : comparison of different study periods.

Figure 4.6 : Example 4.7 – Degree-of-deviation : Bayesian vs. conventional methods.

Figure 4.7 : Probability density and distribution functions for a Gamma distribution with $\alpha = 1$ and $\beta = 3.4$.

Figure 5.1 : The effect of segment length on $P(m_i > R_a)$.

Figure 5.2 : A-False negatives, B – True positives and C – False positives.

Figure 5.3 : Example 5.11 – Degree of *true positive* identifications.

Figure 5.4 : Example 5.11 – Degree of *false positive* identifications.

Figure 6.1 : Illustration of the *regression-to-mean*, *treatment* and *trend* effects.

Figure 6.2 : Illustration of regression-to-mean effect.

Figure 6.3 : Illustration of trend effect.

Figure 7.1 : Case Study 7.1 – Pattern 6 accident freq. vs. left turning flows.

Figure 7.2 : Case Study 7.1 – Pattern 6 accident freq. vs. straight thru' flows.

Figure 7.3 : Case Study 7.1 – Intersection accident patterns.

Figure A1.1 : Appendix A1 – Fitting of Gamma distribution.

CHAPTER 1

INTRODUCTION

1.1 THE ROAD SAFETY SITUATION IN SOUTH AFRICA

The number of accidents and casualties reported in South Africa in 1998 are shown in Table 1.1.

Table 1.1 : 1998 South African accident statistics

Degree	Total	Fatal	Serious	Slight	Damage
Accidents	511605	7260	21265	52097	430983
Casualties	129672	9068	36246	84358	-

Source : CSS Report No. 71-61-01 (1998)

Using the unit cost of accidents compiled by Schutte (2000), as shown Table 1.2, road traffic accidents cost the country approximately R 24817 * 511605 = R 13.7 billion per year (1998 Rands). This figure is about 2 % of the GDP.

Table 1.2 : Unit cost of accidents by severity and status (1998 Rand/accident)

Accident severity	STATUS		
	Drivers and Passengers	Pedestrians	All
Fatal	572386	187562	388487
Serious	122415	49189	88248
Slight	32793	6455	23723
Damage Only	15936	983	15694
Average	24163	32897	24817

Source : CSIR Report CR-2000/4 (2000)

1.2 ACCIDENT CAUSATION

According to Austroads (1994) there are three factors that contribute to motor vehicle accidents :

- Human factors.

- Road environment factors.
- Vehicle factors.

Accidents are often caused not by a single factor but by the interaction of two or more of these factors. Poor driving behaviour in a good vehicle on a good road will in all likelihood present less of a risk than poor driving behaviour in a poor quality vehicle on a poor road.

Research (Austroads ; 1994) has established the contribution of these factors to accidents to be as shown in Figure 1.1.

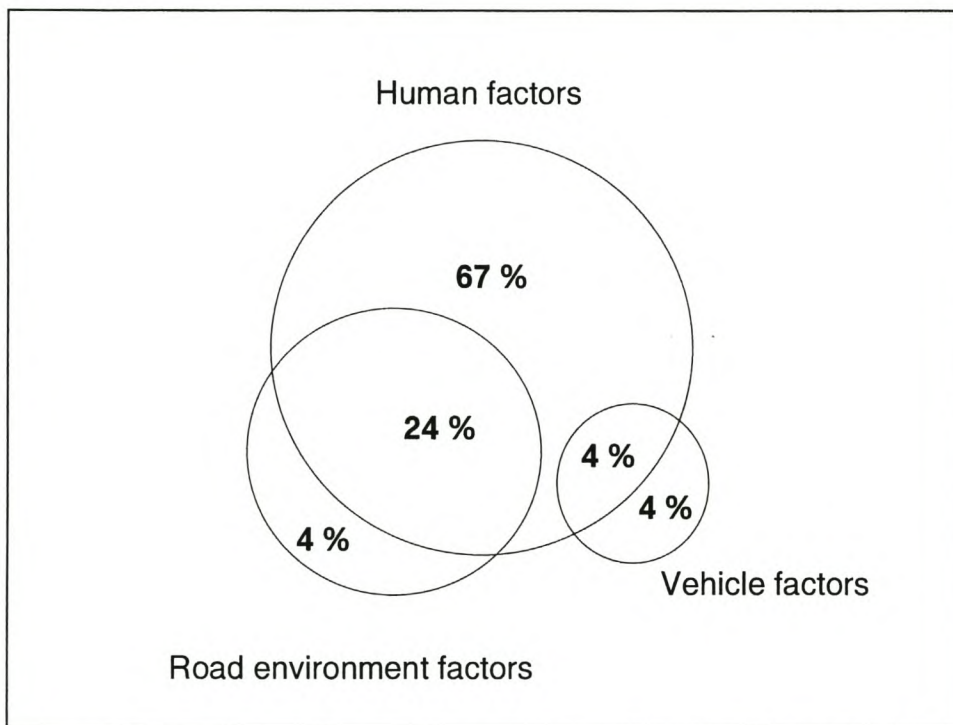


Figure 1.1 : Factors contributing to road traffic accidents (Austroads ; 1994)

From Figure 1.1 it is evident that the road environment contributes either direct or indirectly to approximately 28 % of all accidents. In monetary terms this amounts to about R 3.6 billion per year (1998 Rands).

1.3 ROAD SAFETY MANAGEMENT

The philosophy of road safety management in South Africa centres around the multi-disciplinary approach. This approach advocates that transportation safety can only be addressed through the integrated efforts of the Enforcement, Education and Engineering disciplines.

The primary focus of the Enforcement and Education disciplines is to change road user behaviour in a manner that will lead to an improvement in safety. These disciplines thus focus on those 92 % of accidents that are attributable directly or indirectly to human factors. Typically, enforcement and education campaigns are targeted at high risk behaviours, such as speeding and driving under the influence of alcohol, and high risk road user groups such as pedestrians, taxi drivers, children etc.

The Engineering discipline also plays a vital role in influencing driver behaviour, as engineering measures can influence road user perceptions and ultimately the way road users behave. One of the primary objectives of the Engineering discipline is to make the road environment safe to use, taking into consideration the nature of the interaction between road users and the environment. The environment can be designed and/or modified to accommodate the road user and its limitations and to reduce the severity of accidents should they happen.

According to the *United States Department of Transportation's Strategic Plan : 2000 – 2005*, supporting economic growth is one of the most basic purposes of a national transportation system. Transportation makes possible the movement of people and goods, fuelling the economy and improving the quality of life. However, at the same time the transportation system expose people and property to the risk of accidents and harm. The objective of any road safety management strategy should be to reduce transportation risk, and where possible, to enhance mobility in order to maximise the benefits that can be obtained from a transportation system.

1.4 THE ROLE OF ROAD SAFETY ENGINEERING

The Engineering discipline can contribute to improving the *efficiency* of a transportation system in a number of ways. These can be divided into two categories : a) proactive measures, and b) reactive measures.

a) Pro-active approach

One of the main aims of the proactive approach is to 'build' safety into all aspects of the transportation system. This approach requires a good understanding of the safety implications of engineering decisions relating to the planning, design, implementation, operation and maintenance of road infrastructure elements. In South Africa the Road Safety Manual (COLTO ; 1999) has been developed to facilitate this proactive approach to road safety management. The Manual consists of the following volumes :

- Volume 1 : Principles and Policies
- Volume 2 : Road Safety Engineering Assessments on Rural Roads
- Volume 3 : Road Safety Engineering Assessments on Urban Roads
- Volume 4 : Road Safety Audits
- Volume 5 : Remedial Measures and Evaluation
- Volume 6 : Roadside Hazard Management
- Volume 7 : Design for Safety

The objectives of Volumes 2 and 3 are to provide formal procedures to examine the quality of traffic flow, accident potential and safety performance of a road based on a set number of key indicators to identify hazardous locations and safety deficiencies. A road safety assessment of a road network would, amongst others, identify those entities where the accident potential is high.

The next proactive step is to subject these locations to detailed *road safety audits*, with a view of eventually compiling and implementing a number of remedial measure reports. Volume 4 of the Manual provides detailed guidelines and checklists on how to perform these audits. Not only can road safety audits be performed on existing road safety infrastructure elements, but also on design projects in various stages of execution. According to Volume 4 a Road Safety Audit can be performed during any of the following six stages :

- Stage 1 : Preliminary Stage
- Stage 2 : Draft Design Stage
- Stage 3 : Detailed design Stage
- Stage 4 : During the Construction Stage
- Stage 5 : Pre-opening Stage
- Stage 6 : Existing Road Projects

A Road Safety Audit can be defined as follows (COLTO ; 1999) :

“A Road Safety Audit is a formal examination of a future or existing road/traffic project/any project where interaction with road users takes place, in which an independent, qualified examination team reports on the accident potential and safety performance of the project.”

Volume 6 provides guidelines on how to proactively manage roadside hazards, while Volume 7 provides guidelines on the safety implications of geometric design decisions.

b) Reactive approach

Whereas the focus of the *proactive* approach is on accident potential the focus of the *reactive* approach is on actual accident experience. The assessment procedures of Volume 2 and 3 also consider accident

experience and severity but this is done alongside many other indicators that serve as a measure of accident potential.

The aim of the *reactive* approach is to identify those road infrastructure elements which already operate at unacceptable levels of safety, to investigate these and to apply remedial measures to improve safety.

As far as Road Safety Assessments are concerned the ideal is for an authority to assess their whole road network on a regular basis in order to identify those roads with a high accident potential and to apply preventative measures. The full application of the assessment procedures as described in Volume 2 and 3 of the South African Road Safety Manual (COLTO ; 1999) could be very labour intensive and therefore expensive to implement, especially to road authorities without proper network management systems. It is a South African reality that road authorities have limited budgets for road safety studies and improvements. In light of this reality it would be inappropriate for road authorities to spent money on assessing roads with possibly low accident potentials, while existing hazardous locations with poor safety records continue to operate because of a lack of funds.

Thus in an environment where there are financial and resource constraints it is advisable to first consider *accident experience* only, by implementing road safety remedial programmes, and then if resources allow it, to consider *accident potential* by conducting Road Safety Assessments.

1.5 ROAD SAFETY REMEDIAL PROGRAMMES

The implementation of a road safety remedial programme by a road authority is a very important strategy to achieve a sustained reduction in accidents and severity, and to improve the overall *efficiency* of the road transportation system.

A road safety remedial programme is a process which consists of the following activities :

- a) The identification of hazardous locations.
- b) The preliminary ranking of these identified locations for further study.
- c) Detailed engineering investigation of hazardous locations.
- d) The identification of suitable remedial measures.
- e) Economic evaluation of remedial measure options.
- f) Final ranking for implementation.
- g) Implementation.
- h) Monitoring and evaluation.

a) Identification of hazardous locations

This activity involves the statistical analysis of accident data in combination with road network data and traffic flow information to identify those locations which experience an abnormally poor level of safety when compared to similar locations.

b) Preliminary ranking of hazardous locations

The amount of financial, human and physical resources available to a road authority might be such that it is unable to conduct detailed investigations of all the identified hazardous locations. To ensure the efficient allocation of resources is it important to rank sites according to their expected economic benefit and then to apply resources in descending order of priority.

c) Detailed investigation

In order to identify the most effective and efficient remedial measures it is important to have a thorough understanding of the extent, nature and causes of the accident problem at a location.

An investigation could be one or a combination of the following :

- A further detailed analysis of accident data to reveal accident types and patterns that fall outside the 'norm'. This information could provide clues as to the causes of accidents and casualties.
- Detailed analysis of individual accident records – accident reconstruction and analysis.
- A formal Stage 6 Road Safety Audit according to the guidelines of Volume 4 of the South African Road Safety Manual (COLTO ; 1999).
- Conflict studies and analysis

d) The identification of suitable remedial measures

Once all the contributory factors to the safety problem/s have been identified the next step is to identify appropriate remedial measures. *Volume 5 : Remedial Measures and Evaluation* of the South African Road Safety Manual (COLTO ; 1999) provides guidance to the road safety engineer on choosing appropriate remedial measures to address particular problems.

e) Economic evaluation of remedial measures

To ensure that resources are allocated to those projects that will yield the best economic returns is it necessary to conduct an engineering economic study to determine the expected economic return of an remedial measure 'investment'. *Volume 5 : Remedial Measures and Evaluation* of the South African Road Safety Manual (COLTO ; 1999) provides details on expected accident reductions associated with different remedial measures and proceeds to show how the expected

Net Present Value, Benefit/Cost ratio and Internal Rate of Return can be estimated.

f) Final ranking and selection for implementation

Once the expected cost and economic returns of each hazardous location have been determined a decision has to be made on which locations to select for the implementation of remedial measures. The number of sites finally selected for treatment will depend on the available budget. Selecting sites can be a complicated exercise which falls in the realm of transport economics. Some of these selection methods are discussed in Volume 5 : Remedial Measures and Evaluation of the South African Road Safety Manual (COLTO ; 1999).

g) Implementation

The identified remedial measures are then designed and implemented at those locations selected during the previous step.

h) Monitoring and evaluation

After implementation is it imperative that the *effectiveness* and *efficiency* of the remedial measures implemented be evaluated.

The focus of an evaluation study should be on estimating the degree of change in the level of safety and the economic benefit associated with this change.

This information is required to :

- a) Ensure public accountability with regards to the spending of public funds.



- b) Add to the database of knowledge on the effects of different types of remedial measures in order to provide better quality information for future road safety studies.

Apart from just evaluating the change in the level of safety it could also be necessary to evaluate the impact of a remedial measure on social issues, environmental issues, traffic flow operations, land use and security issues.

1.6 THE ROLE OF ACCIDENT DATA ANALYSIS

Three of the steps of an road safety remedial programme rely exclusively on the analysis of accident data – the identification and preliminary ranking of hazardous locations and the evaluation of road safety remedial measures. The analysis of accident data can also assist in identifying accident causes and appropriate remedial measures, however, it is possible (although not recommended) to perform these steps without analysing accident data.

According to Hauer (1997) accidents are the physical manifestation of 'unsafety'. The proper analysis of accident data is therefore the most appropriate way to gain an understanding of road safety and all its dimensions.

When identifying hazardous locations it is important to identify those locations that are truly hazardous and to 'miss' those locations which are not truly hazardous. Not identifying 'true' hazardous locations, could cause these hazardous locations to remain untreated, obviously with potentially severe consequences in terms of deaths, injury and damage to property. Identifying 'false' hazardous locations could waste potentially scarce resources to investigate locations which are not really unsafe and whose potential economic returns are low.

For the sake of the overall efficiency of a road safety remedial programme is it important that the analysis of accident data to identify hazardous locations

use methodologies that are *efficient* – i.e. methods that maximise the degree of ‘true’ identifications and minimise the degree of ‘false’ identifications.

It is important to use accident data analysis methodologies that will produce accurate and reliable estimates of the safety effect. The *underestimation* of the safety effect could cause a treatment to be discarded in favour treatments whose ‘true’ effects are less. It could also cause the expected economic returns at a location to be *underestimated*, with the possible consequence that a perfect viable location remains untreated. Similar principles apply when the safety effect is *overestimated*.

1.7 STUDY OBJECTIVES

The objective of this thesis is to report on, investigate and present suitable accident data analysis methodologies for the efficient identification and ranking of hazardous locations and the estimation of accurate safety effects of road safety engineering remedial measures.

Issues relating to the detailed investigation of hazardous locations, the identification of remedial measures and the economic evaluation of remedial measures will not be addressed in this thesis.

Chapter 2 will investigate issues relating to the information required to efficiently identify and rank hazardous locations and to accurately estimate a treatment effect. Particular attention will be paid to issues affecting the accuracy and management of data.

Chapter 3 deals with general aspects concerning the measurement of safety. The concept of *measurement error* will be explained and steps to reduce the error associated with a measurement will be presented. The conventional methods (i.e. accident rates) to express the risk associated with the road transportation system will be critically evaluated and certain remedial strategies e.g. using Safety Performance Functions, will be proposed to overcome certain inherent shortcomings of the conventional approach.

Chapter 4 deals with the actual accident data methodologies that can be used to obtain accurate estimates of the level of safety for an single entity or a group of entities. The two approaches to road safety measurement, the *Conventional* and *Bayesian* approaches will be presented. The performance of these methods will be assessed by means of experiments, the details of which are discussed and presented in Appendix A1 and A2. The use of the methods will also be illustrated using accident data and traffic flow data on Class 1 2-lane rural roads in the Province of the Western Cape (Appendix B1).

Chapter 5 will present and evaluate the different Conventional and Bayesian methodologies available to identify and rank hazardous locations. The performance of the different methods will also be assessed by means of a number of experiments, the details of which are contained in Appendix A3. The use of the methods will also be illustrated using accident data and traffic flow data on Class 1 rural roads in the Province of the Western Cape (Appendix B2 and B3).

Chapter 6 will present and evaluate the different Conventional and Bayesian methodologies to evaluate the effectiveness of road safety remedial measures. Chapter 6 is largely based on the work of Dr Ezra Hauer as contained in his authoritative book on the subject – *'Observational Before-and-After Studies in Road Safety'* (Pergamon ; 1997). Once again the different methods will be assess using a series of experiments, the details of which are contained in Appendix A4.

Throughout Chapters 3, 4, 5 and 6 it will become evident that multivariate regression models play a very important role in the Bayesian estimation of safety, identification and ranking of hazardous locations and the evaluation of road safety remedial measures. Chapter 6 provides a general overview of issues relating to the modelling of accident data. The objective is not to make the reader proficient in developing multivariate regression models, but to provide general information on how these models are developed and applied.

CHAPTER 2

INFORMATION REQUIREMENTS

2.1 INTRODUCTION

Apart from information on accident frequencies and severities, the successful implementation of a road safety remedial programme also require information on the characteristics of the road network and information on traffic flows. This information, amongst others, may be required to calculate accident rates, to identify suitable reference and comparison groups and to develop multivariate regression models.

The objective of this Chapter is to discuss issues relating to the quality and management of these information sources – accident data, road network information and traffic flow information.

Good quality accident data is absolutely essential for the efficient implementation of a road safety remedial programme. A number of issues affecting the quality of accident data in South Africa will be reported on. One of the main issues that could compromise the quality of accident data, namely the underreporting of accidents will be discussed in detail. It will be shown that the underreporting of accidents is a widespread which could seriously compromise the efficiency of a road safety remedial programme. A methodology will be presented to quantify the effect of underreporting on road safety measurement and evaluations.

2.2 ACCIDENT DATA

2.2.1 DATA QUALITY

According to O'Day (1993) for accident data the components of quality include:

- Completeness of coverage – the degree to which the data management system contain all the accidents as defined by legislation.
- Consistency of coverage – where the degree of reporting varies geographically or by time, weather or other factors.
- Missing data – the degree to which there are missing data elements for those accident records that are reported.
- Consistency of interpretation – whether the report elements, for example the degree of injury, are reported consistently by all persons that investigate and report on accidents.
- The correct data – whether the correct data are being collected at the appropriate level of detail.
- Correct data capturing procedure - whether the data as they appear on the accident report form are taken up correctly and without error in a computerised database.

2.2.2 REPORTABILITY

The South African legislated definition of an accident is contained in Paragraph 61(1) of the National Road Traffic Act 93 of 1996 :

“ The driver of a vehicle on a public road at the time when such vehicle is involved in or contributes to any accident in which any other person is killed or injured or suffers damage in respect of any property or animal shall...” (Italics added)

In Paragraph 61(1f) it is stated :

“...unless he or she is incapable of doing so by reason of injuries sustained by him or her in the accident, as soon as reasonably practicable, and in any case within 24 hours after the occurrence of such accident, report the accident to any police officer at a police station or at any office set aside by a competent authority for use by a traffic officer....”(Italics added)

The Act defines a ‘public road’ as follows :

‘Public road’ means any road, street or thoroughfare or any other place (whether a thoroughfare or not) which is commonly used by the public or any section thereof or to which the public or any section thereof has a right of access. And includes –

- a) The verge of any such road, street or thoroughfare ;
- b) Any bridge, ferry or drift traversed by any such road, street or thoroughfare; and
- c) Any other work or object forming part of or connected with or belonging to such road, street or thoroughfare; (Italics added)

The Act defines a ‘driver’ as follows :

‘Driver’ means any person who drives or attempts to drive any vehicle or who rides or attempts to ride any pedal cycle or who leads any draught, pack or saddle animal or herd or flock of animals, and ‘drive’ or any like word has a corresponding meaning. (Italics added)

Act 93 of 1996 does not specify a definition of the concept ‘damage’. In other words in terms of the Act, for example, two vehicles that collide without any visible damage and injury to any person would not constitute an accident. A chipped windscreen as a result of a loose stone on the road would constitute an accident so would hitting an animal even if it caused no damage to the

vehicle. The event of a passenger falling inside a bus because the bus braked too sharply would constitute an accident in terms of the Act.

In some countries, such as Canada, an accident is only reportable if the damage to the vehicle exceeds a minimum amount (Hauer ; 1997). In 1990 the reportability limit in Ontario was \$700. A reportability limit has its drawbacks in the sense that the cost to fix damage to a car could change from area to area, and secondly if the threshold is not continually adjusted for inflation it could cause that more and more damage only accidents are becoming reportable.

According to Hauer (1997) many countries only keep records of injury accidents. The injury accident count does not depend on the cost of car repairs or the value of money.

2.2.3 ACCIDENT CLASSIFICATION

According to the Opperman and Hutton (1991) in South Africa a road traffic accident can be classified into one of four categories :

- **Fatal accident**

An accident that results in injuries that cause immediate death, or death within 6 days as a direct result of the accident.

- **Serious injury accident**

An accident that results in injuries that include fractures, concussions, severe cuts and lacerations, shock necessitating medical treatment and any other injury that requires hospitalisation or confinement to bed.

- **Slight injury accident**

An accident that results in injuries that include cuts and bruises, sprains and slight shock not requiring hospital treatment.

- **Damage Only accident**

An accident in which there is no personal injury but damage to property.

According to Lötter (2000) up to 1975 the official definition of a fatal accident in South Africa was one where death occurred within three months of an accident. From January 1975 the definition was changed to death within 6 days of an accident.

Lötter (2000) notes that in South Africa there are no formal follow-up procedures for tracking the progress of traffic accident casualties en route to hospitals⁶ or while undergoing hospital care. The accident report forms are therefore not updated as far as fatalities are concerned. Consequently fatality statistics mostly reflect 'dead on the scene' cases. The police reported fatalities are therefore heavily underreported when compared to actual fatalities (according to the 6 day definition).

Lötter (2000) found that approximately 76.4 % of fatalities occur at 0 days after the accident while 14.7 % occur within 1 – 6 days after the accident. It can therefore be concluded that fatalities could be underreported by as much as 23.6 %.

2.2.4 ACCIDENT REPORTING

In South Africa the responsible party for the investigation of accidents and the completion of the accident report form is the South African Police Services. (SAPS). The law does however provide for duly authorised traffic officers to complete accident report forms. The traffic departments of certain towns such as for example, Stellenbosch in the Western Cape, has established their own traffic accident units which have taken over the function of accident investigation and reporting from the local SAPS.

Prior to 1999 accident data information was recorded on the SAP352 accident report form. This form was completed by the SAPS in triplicate. The original copy was kept by the SAPS for their own records. The 2nd copy was send to the Central Statistical Services where the information was taken up into the National Accident Database. The 3rd copy was made available to the relevant road authority (if any).

In the event of a fatal accident or if an accident was caused by an serious offence ,the SAPS would open a case docket. As information from the accident report form has to be incorporated into the criminal investigation the 3rd copy was often filed with the case docket without a copy thereof made available to the relevant traffic authority

It is therefore highly likely that fatal accidents in the accident data management systems of local, regional and provincial authorities are under reported.

During 1999 most Provinces changed over to a new accident report form called the OAR (Officer Accident Report) form. Also in 1999 the responsibility to operate and maintain a National Accident Register shifted from Stats SA (Statistics South Africa) to the National Department of Transport. Since there is only 1 copy of the OAR form as opposed to the 3 copies of the SAP352A form, large scale changes in the status quo were necessary.

Firstly, it was expected of all the provinces to establish their own provincial and/or regional databases and to provide accident data electronically to the National Accident Register. The SAPS is still primarily responsible for completing the accident form. Each Province in South Africa has their own strategy for collecting these forms from the SAPS and ensuring that data are taken up into a provincial accident database, and from there to the National Accident Register. In the case where the SAPS has to keep the original OAR form for their case docket they are obliged to make a photocopy of the form available for collection within 7 days after the accident. In the Western Cape mechanisms are in place to ensure that all accident forms are collected and that forms don't 'disappear' into case dockets as was the case with the SAP352 forms. It is therefore likely that in some instances the number of fatal accidents will appear to increase as a result of better reporting.

2.2.4.1 UNDERREPORTING

According to Hauer and Hakkert (1988) much of what we know and do about road safety is tied to the use of accident data reported to and by the Police.

For instance if the level of accident reporting was to decrease the ability to manage road safety will be compromised. Underreporting will result in fewer accidents being reported with the consequence that it will take longer to accumulate the same amount of data. Hazardous locations will take longer to detect, accident patterns will be more difficult to discern, the effect of safety remedial measures will be less precisely known etc. (Hauer and Hakkert ; 1988).

According to a study done by Hauer and Hakkert (1988) amongst 18 reporting authorities in North America (USA and Canada), Europe (Netherlands, Germany) the degree of accident underreporting is substantial and that it differs widely from one authority to another. They estimated that fatalities seem to be known to an accuracy of ± 5 %. It was also found that 20 % of injuries that require hospitalisation are underreported and only about 50 % of all injuries sustained in motor vehicle accidents are reported to the police.

In a detailed comparison of accident reporting levels in 13 different countries Elvik and Mysen (1999) found that reporting levels varied widely between different countries and that reporting were incomplete at all levels of injury severity. They found the mean reporting level for fatalities to be about 95 % (according to the 30 day rule), for serious injuries 70 % (require hospitalisation), slight injuries 25 % and very slight injuries 10 %.

According to Hauer and hakkert (1988) and James (1991), for injury accidents the age of the casualty, type and number of vehicles involved in the accident, accident location, severity of injury, and consequently mode of transport to

medical care and/or length of time before treatment, affected whether or not an accident was likely to be reported to the police.

It is evident from the studies by Hauer and Hakkert (1988), James (1991) and Elvik and Mysen (1999) that fatal accidents are reported more fully than serious accidents and that the reporting of the latter is better than that of slight injury accidents.

Hauer and Hakkert (1988) states that the probability of reporting an injury young children is 20 – 30 % and for people over 60 it is 70 %. James (1991) explains that the low reporting rate for children is related to the type of accident in which they are most likely to be involved in, namely bicycle accidents.

Hauer and Hakkert (1988) found that that the probability of reporting an injury is largest for the driver, less for the passenger and even less for non-occupants. This was confirmed by Elvik and Mysen (1999) who found that reporting levels tend to be higher for occupants and lowest for cyclist, and that this was the pattern for all 13 countries considered in their investigation. They found that the reporting of single-vehicle bicycle accidents is particularly low – below 10 % in all the countries studied.

Smith (in Hauer and Hakkert ; 1988) found that 57 % and 12 % reporting levels for single-vehicle and damage only accidents respectively ; for multi-vehicle accidents the corresponding percentages are 96 and 41.

Hauer and Hakkert (1988) argues that most of what is said about road safety is based on accidents that have been reported and not on estimates of what actually occurred. Not only do such statements make the safety problem appear to be smaller than it really is, they also mix and confuse changes and trends in safety with changes in trends in the inclination to report accidents.

According to Hauer and Hakkert (1988) if the inclination to report an accident is constant from time period to time period and between sites, comparisons on

safety on the basis of reported accidents are legitimate. They do however argue that this assumption is unrealistic. There are a number of factors that influence the probability of an accident being reported – factors that can change over time and from location to location.

Thus if the probability of reporting an accident is not constant across time and space accurate assessments of safety cannot be made without knowing what the probabilities are. Without knowing the degree of underreporting a reliable estimate of the ‘true’ level of safety at a location cannot be determined.

Hauer and Hakkert (1988) provided the following methodology to assess the impact of underreporting on the estimation of safety :

Let :-

X_i - The number of accidents of class i reported to the police.

p_i - The probability that an accident of class i will be reported to the police.

m_i - The actual number of accidents of class i expected to occur at the site.

The expected number of reported accidents is given by r_i :

$$r_i = p_i m_i \quad \dots[2.1]$$

Therefore

$$m_i = r_i / p_i \quad \dots[2.2]$$

The value of p_i is an estimate and is surrounded by uncertainty. Two scenarios will now be investigated, a) p_i is known exactly, and b) p_i is uncertain.

a) The accuracy of the estimate of m_i is described by $VAR(m_i)$:

$$VAR(m_i) = \frac{VAR(r_i)}{p_i^2} \quad \dots[2.3]$$

When the Poisson model is applied to accident data, the variance is equal to the mean, but when accident reporting is not complete the variance of the estimate will always be larger than the mean even if p_i is known precisely (Hauer and Hakkert ; 1988).

If it is assumed that there are n annual accident counts for some entity and m_i is the expected number of reported accidents per annum then :

$$VAR(r_i) = \frac{r_i}{n} \quad \dots[2.4]$$

Therefore

$$VAR(r_i) = \frac{m_i p_i}{n} \quad \text{and} \quad \dots[2.5]$$

$$VAR(m_i) = \frac{m_i}{np_i} \quad \dots[2.6]$$

The *variance-to-mean* ratio, which serves as a measure of accuracy, is given by :

$$\frac{VAR(m_i)}{m_i} = \frac{1}{np_i} \quad \dots[2.7]$$

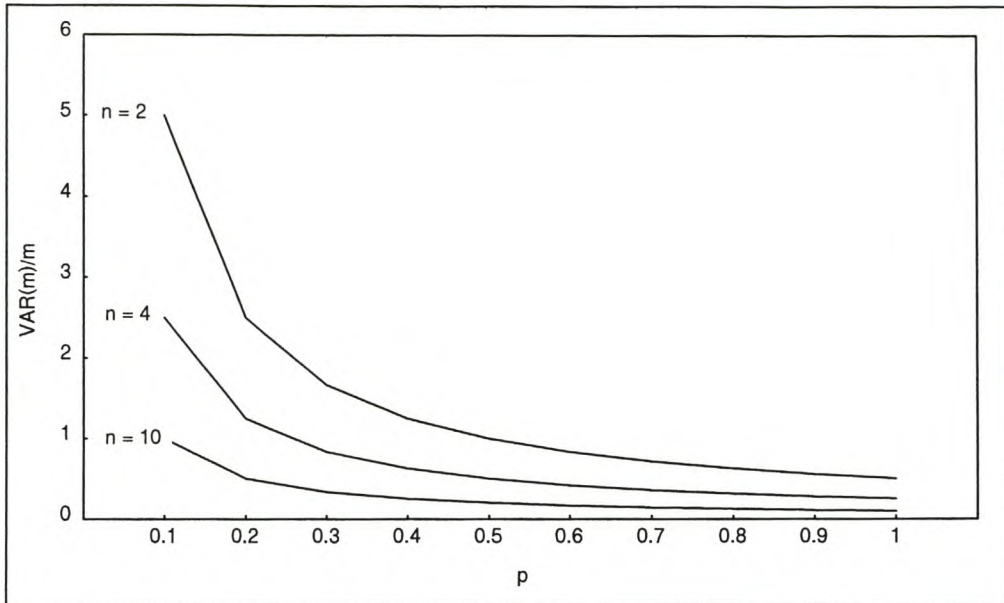


Figure 2.1 : Accuracy vs. degree of accident reporting

From Figure 2.1 it is evident that if accident reporting is complete (i.e. $p = 1$) only 2 years of accident data is required to get a variance that is half its mean. To keep the same level of accuracy with a 50 % reporting level 4 years of accident data is required.

- b) If the reporting probability is known only with some uncertainty i.e. it is a random variable, it can be shown (Hauer and Hakkert ; 1988) that the variance of the expected accident frequency is given by :

$$VAR(m_i) = \frac{m_i}{np_i} + \frac{m_i^2 VAR(p_i)}{p_i^2} \quad \dots[2.8]$$

It is evident that the number of years for which accident counts are available (n) affects only one component of the variance of m_i and not the other. Thus for no matter how many years of accident counts are reported, the uncertainty surrounding p_i puts a limit on how accurately m_i can be measured.

If the goal is to obtain estimates of the ratios of the m_i 's 'before' and the m_i 's 'after' treatment the probabilities of reporting during these two periods should also be considered. If p_i is the same for these two periods then the net effect of incomplete accident reporting is merely to reduce the amount of accidents that can be collected per unit time.

Hauer and Hakkert (1988) argue that it is illogical to assume that the p_i in the 'before' and the 'after' periods will be the same because remedial action may change the accident patterns with some type of accidents having a better probability of being reported than others.

If it can be assumed (however questionable) that p_i is the same for the 'before' and 'after' period the accuracy of the 'after' and 'before' ratio θ_i is given by (Hauer and Hakkert ; 1988):

$$VAR(\theta_i) = \theta_i^2 \left[\frac{1}{r\theta_i} + \frac{1}{r_i} + \frac{2VAR(p_i)}{(p_i)^2} \right] \quad \dots[2.9]$$

EXAMPLE 2.1

Hauer and Hakkert (1988) provide the following example :

Suppose the number of reported injury accidents changed from 25 before treatment to 20 after treatment then $\theta = 20/25 = 0.8$ and ARF (Accident Reduction Factor) = 20%.

If p is known exactly and is equal between the 'before' and 'after' periods $VAR(\theta)$ is given by :

$$VAR(\theta) = (0.8)^2 [1/20 + 1/25] = 0.058$$

If however if p is a random variable with $E(p) = 0.7$ and $VAR(p) = 0.01$ the value of $VAR(\theta)$ is given by Equation 2.9:

$$VAR(\theta) = (0.8)^2 [1/20 + 1/25 + 2(0.01)/(0.7)^2] = 0.083$$

2.2.4.2 INCOMPLETE/INCORRECT REPORTING

The problem of non-reporting is compounded by a variety of inaccuracies and errors recorded by the police on the accident report form. Hauer and Hakkert (1988) list the following common problems

a) The police often misclassify the severity of injuries. James (1991) reports on research that estimated that the net effect of police misclassification caused the number of seriously injury cases to be under-reported by 13 %. The problem of misclassification was also identified in Sweden by Thorson and Sande (in James ; 1991) who found that 21 % of in-patients were wrongly classified by the police as slightly injured.

b) Data fields are left uncompleted.

Depending on the nature of an accident study uncompleted data fields could potentially introduce a serious bias in the analysis. Take for example the *Quality of road surface* field on the OAR form. The person completing the OAR form might at all times be able to identify when a road surface is *good*, but when it is not good he/she might not be able to choose from the remaining options – *bumpy, pothole, cracks, corrugated, other* – and rather choose nothing at all and leave the field blank. Omitting all records in this case which have blank fields will introduce a bias in favour of '*good*'.

c) Very imprecise location of an accident is given, especially if the accident did not occur at an intersection.

Having accurate and reliable information on the location of an accident is a prerequisite for using accident data to identify hazardous locations. The Provincial Accident Data Centre which collect OAR forms from more than 120 SAPS stations in the Western Cape are experiencing

tremendous problems because of the poor quality of the accident location description. Problems are especially experienced with *Damage Only* type accidents when the driver/s reported the accident directly at the police station.

The accuracy of the locational description on the accident report form is a function of the level of motivation of police officers and the importance that they place on the accuracy of this data item. There appears to be, in the Western Cape at least, a lack of awareness of the importance of this data item to overall road safety management. The South African Police Services are generally understaffed and consider accident reporting to be of less importance than other policing duties such as crime prevention.

2.3 ROAD NETWORK INFORMATION

The analysis of accident data often requires the comparison of safety between different locations, e.g. during the identification of hazardous locations. According to Hauer (1995) accident rates/number can only be compared if there exists a reasonable expectation that the accident rate between two locations should ideally be the same i.e. there exist a reasonable *expectancy of equality*.

A major factor that could cause the level of safety between two locations to be different is the geometric design characteristics of the locations in question.

Bester (1994) found that the following geometric design factors have a significant influence on the safety of rural roads : Number of lanes, lane widths, shoulder widths, terrain types, riding quality and the type of shoulder (paved or unpaved). In addition to the factors identified by Bester (1994), Milton and Mannering (1998) identified the following additional geometric factors ; vertical grades, speed limits, horizontal curve radii, horizontal curve central angles and tangent lengths.

The safety of intersections could be a function of the type of intersection control, traffic signal phasing and timing design, the number of approach lanes, the presence of median islands, the width of such median islands, the presence of pedestrian crossings, approach speeds etc.

Other factors that could influence the level of safety is the climate as well as whether the road is situated in an urban, rural or semi-rural area.

The *Empirical Bayesian* approach and some of the *Conventional* methods, particularly to identify hazardous locations, require a reference group to obtain an estimate of the safety at a location. Such a reference group should consist of sites that share similar geometric and environmental characteristics to the site/s under investigation.

Conventional evaluation methods often require an comparison group to account for changes in traffic volumes and other external influences over time. It is possible that the degree and extent of changes in traffic volumes and other external influences could depend on the geometric and environmental characteristics of the site in question. In such a case it would be desirable to have a comparison group of sites that share similar geometric and environmental characteristics. In order to eliminate the *regression-to-mean* effect it is necessary that the comparison group of sites not only share similar geometric and environmental characteristics but also similar levels of safety as the site/sites in question.

In conclusion therefore, to implement a road safety remedial programme sufficient information should be available on the geometric and environmental characteristics of the whole road network.

2.4 TRAFFIC VOLUME INFORMATION

Another factor that could cause accident rates/numbers between locations to be different is the traffic volume.

Traffic volumes are used to estimate exposure in accident rate calculations. It is also an important input variable (often the only variable) when developing accident models (Safety Performance Functions) using regression techniques.

Traffic volumes, especially on rural road networks, are often expressed as an AADT – Annual Average Daily Traffic. According to Papenfus (1992) the AADT can be defined as that traffic volume which, if multiplied by the number of days in a year, will yield the total annual traffic volume on a road.

The extent of a typical rural or urban network makes it impractical for each link and node to be counted continuously in order to obtain an exact AADT value. In practice, therefore, traffic counts are collected on a sampling basis in which counts are collected only for short periods of a year. The estimation of AADT's from these short term counts and not from a continuous 365 day count have implications regarding the accuracy of the resultant AADT and accident rate estimates.

2.4.1 ACCURACY OF TRAFFIC COUNTS

The accuracy of an AADT estimate depends on a number of factors :

- The short term counting period.
- The day /s of the week on which counting was conducted.
- The stratification method.

According to Papenfus (1992) the stratification method is based on the assumption that the different traffic patterns which occur on the links of a road network can be divided into different groups or strata. Permanent counting stations are installed per stratum to determine average expansion factors for each stratum. These expansion factors are used to convert the observed short term count to an AADT count.

The stratification method refers to how strata are defined and how the different links are divided into the different strata.

a) No stratification.

Only one set of expansion factors is used for all roads.

b) Elementary stratification.

Only peak hour traffic is used for stratification.

c) Night-time traffic stratification.

The percentage traffic before 06:00 and 18:00 is used to estimate weekend traffic and for stratification purposes.

d) Full stratification .

Stratification is based on both the peak hour traffic and the estimated weekend traffic.

e) Mother/daughter method.

According to Papenfus (1992) this method is based on the assumption that for each short term or “daughter” counting station, a similar traffic pattern exists at a permanent or “mother” counting station. Expansion factors are calculated for the mother station and applied at the daughter station for the estimation of the AADT.

f) Direct estimation.

According to Papenfus (1992) if traffic counts are made over a period of seven normal days or more, the AADT can be estimated directly using the ‘mother/daughter’ method without making use of stratification.

Papenfus (1992) evaluated the accuracy of different stratification methods for different collection periods. The coefficient-of-variation associated with AADT estimates based on different counting periods and stratification methods are shown in Table 2.1 and Table 2.2.

The data in Table 2.1 is only applicable to links where the AADT > 500. It is assumed that a 12 hour count is from 06:00 to 18:00, a 18 hour count from 04:00 to 22:00 and a 24 hour count from 00:00 to 24:00.

Table 2:1 : Coefficient-of-variation for different stratification methods and counting periods

Days counted	Hours counted per day	None	Elementary	Night-time	Full
MON	12 hr	14.18	12.88	12.88	10.95
	18 hr	12.39	11.2	10.67	10.21
	24 hr	11.32	10.62	10.01	9.88
TUE	12 hr	15.83	14.81	14.81	11.54
	18 hr	13.72	12.6	12.21	10.54
	24 hr	13.04	12.22	11.96	10.56
WED	12 hr	15.83	15.04	15.04	11.15
	18 hr	13.67	12.79	12.31	10.11
	24 hr	12.92	12.29	11.75	9.92
THU	12 hr	14.51	13.82	13.82	10.04
	18 hr	12.09	11.35	10.86	9.04
	24 hr	11.32	10.84	10.15	8.83
FRI	12 hr	12.19	11.83	11.83	10.92
	18 hr	9.74	9.83	9.9	9.65
	24 hr	9.83	9.9	9.74	9.57
MON - TUE	12 hr	14.17	12.97	12.97	9.96
	18 hr	11.61	10.31	9.75	8.6
	24 hr	10.41	11.14	8.86	8.23
TUE -WED	12 hr	15.3	14.34	14.34	10.2
	18 hr	12.87	11.77	11.16	8.83
	24 hr	11.95	11.14	10.26	8.5
WED - THU	12 hr	14.76	13.99	13.99	9.88
	18 hr	12.31	11.45	10.91	8.7
	24 hr	11.46	11.14	9.97	8.42
THU - FRI	12 hr	12.18	11.62	11.62	9.44
	18 hr	8.86	8.48	8.41	8.17
	24 hr	8.21	11.14	8.07	8.02
SUN - SAT	12 hr	9.69	9.7	9.7	8.34
	18 hr	7.23	7.14	7.07	7.05
	24 hr	6.51	6.67	6.67	6.84

Table 2.2 : Coefficient-of variation values

Number of weeks counted	Mother/Daughter	Direct Estimation
1	6.91	6.33
2	6.71	6.04
4	5.87	5.15
8	5.06	4.33

Sweet and Lockwood (1983) undertook an investigation into the accuracy of short term traffic counts. They identified 4 main traffic categories :

- a) Urban
- b) Strategic
- c) Recreational
- d) Rural low flow

Each one of these categories is characterised by distinctive types of traffic variation over the year.

a) Group A – Urban

This group have little traffic variation throughout the year, but with slightly lower flows during holiday periods. This group includes both urban commuter and non-commuter traffic. The main difference in the profiles of commuter and non-commuter traffic lies in the weekend traffic volumes. Commuter traffic has a lower volume over the weekend than non-commuter traffic.

b) Group B - Strategic

These are primarily interurban routes with a high proportion of heavy and commercial vehicles. Weekend traffic volumes are high particularly on Saturdays. The AADT can range from 500 to 8000 vehicles per day.

c) Group C - Recreational

This group is characterised by low to medium traffic volumes : 300 to 3000 vehicles per day. The flows are heavily influenced by recreational traffic. A site is classified as recreational if the holiday traffic is on average 50 percent higher than the normal traffic.

d) Group D – Rural low flow

This group comprises rural low flow routes with little or no recreational element, and includes most gravel roads and other minor roads. Daily flows tend to be below 1000 vehicles per day.

Table 2.3 contains *coefficients-of-variation* for these different traffic categories for three different common counting periods.

Table 2.3 : Coefficients-of- variation (%).

Traffic category	Count period		
	1-day	2-day	7-day
URBAN	9.6	7.9	5.5
STRATEGIC	17.7	14.3	9.0
RECREATIONAL	27.1	22.1	15.2
RURAL LOW FLOW	15.8	12.5	7.3

Source : Sweet et al. (1983)

The information in Table 2.3 is based on the assumption that the 1-day and 2-day counts were conducted for a full 24 hours on normal weekdays i.e. not during weekends and holiday periods and that the 7-day count was not conducted during holiday periods.

In conclusion, the reliability and accuracy of road safety estimates that make use of traffic flow information such as AADT's depends directly on the accuracy with which the AADT is estimated. Since the accuracy of an AADT is a function of the collection period, day of the week and stratification method prior knowledge of these variables is required. Should information on the day of the week and stratification method not be available it is recommended that the coefficient-of-variation values in Table 2.3 be used.

2.5 INFORMATION MANAGEMENT

In the preceding sections it was emphasised that in order to implement an efficient and effective accident remedial programme information is required on accident data, road network data and traffic flow data.

The objective of this section is to show how accident, road network and traffic flow data could be managed in order to facilitate the efficient and effective implementation of a road safety remedial programme.

In most road authorities information will be stored in a number of computerised databases. Typically accident data, road network information and traffic count information will be stored in separate databases. Each of these databases are normally operated and maintained by separate management systems which allow for data input, verification, database maintenance, data extraction and reporting.

In order to implement a road safety remedial programme and to analyse accident data according to the procedures set out in this document it is important to be able to combine the information in these databases in a manner that will satisfy the needs of an accident data analysis study.

Section 2.5.1 will provide a brief introduction into database design principles and will discuss the principle of *normalised tables* and how *normalised tables* can be combined to retrieve information. Section 2.5.2 will provide an introduction to SQL (Structured Query Language) and will illustrate by means of numerous examples how SQL can be used to retrieve accident data, network data and traffic flow information from various database tables in a format that is suitable for direct analysis or for further manipulation using other data analysis tools such as spreadsheets.

2.5.1 DATABASE DESIGN

2.5.1.1 BASIC PRINCIPLES

Table 2.4 : Components of database design

Table	In a relational database data are organised in tables. Fig. 2.2 shows an example of a table that holds information on accidents.
Rows (Records)	A table holds information on an event (or entity) in an horizontal row. The table in Fig. 2.2 has 4 records i.e. it holds information on 4 different accidents.
Columns (Fields)	Information on an event (entity) is stored in columns. In Fig. 2.2 each accident record has 8 attributes: an accident identification number, a road identification code, the location of the accident, the date of the accident, the number of fatalities, serious injuries, slight injuries and magisterial area. Each attribute is defined as a column.
Primary Key	A primary key is a column/s that uniquely identify a record. In Fig. 2.2 the 'accnum' column contains information that is unique to each accident and is therefore the primary key.
Field	A field is defined as the intersection of a row and a column. A field may or may not contain data. If there is no data in a field it is said to contain a NULL value.

Table : accDetail					
accNum	Road	AccKm	AccDate	Fatal	MagArea
1	TR00901	12.3	12/11/99	2	Bellville
2	TR00101	34.45	13/10/99	0	George
3	TR01101	14.20	14/11/99	0	Vredendal
4	MR0203	45.45	23/10/99	1	Paarl

Figure 2.2 : A database table

2.5.1.2 THE RELATIONAL DATABASE CONCEPT

It is desirable that data should be stored in *normalised* tables, that is tables that do not contain redundant information. Redundant information is information that is repeated unnecessarily. The use of normalised tables will

lead to less storage space required and that the potential for data errors and inconsistencies are reduced.

accNum	Road	AccKm	AccDate	Fatal	MagArea	Region
1	TR00901	12.3	12/11/99	2	Bellville	Metro
2	TR00101	34.45	13/10/99	0	George	S Cape
3	TR01101	14.20	14/11/99	0	Vredendal	W Coast
4	MR0203	45.45	23/10/99	1	Paarl	Boland

Figure 2.3 : A database table with redundant information

The Table in Fig. 2.2 could be expanded to include an attribute called *Region* – the region in which the accident occurred. This is indicated in Fig 2.3. The Table in Fig.2.3 is not normalised since it contains redundant information – *Region*. Since it is always known in which region a specific magisterial area is situated it is not necessary to include *Region* in the *accDetail* table. This problem can be overcome by creating a new table, *magDetail*, (see Fig. 2.4) which contains information on in which region a magisterial area is situated. In order to determine in which region an accident occurred the Tables in Figures 2.2 and 2.4 can be linked. This concept of linking tables is the topic of the next section.

magArea	Region
Bellville	Metro
Paarl	Boland
Vredendal	W Coast
George	S Cape

Figure 2.4 : A database table

2.5.1.3 RELATIONSHIPS BETWEEN TABLES

Two or more tables in a database can be linked if they have an attribute/s in common.

a) One-to-one relationship

If a record from Table A can only be related to one record from Table B and vice versa there exists a one-to-one relationship between Table A and Table B.

b) One-to-many relationship

If any record from Table A is related to zero or more than one record in Table B and a record in Table B is related to only one record in Table A there is a one-to-many relationship between Tables B and A. In our previous example an accident can only be associated with one magisterial area while a magisterial area can be associated with a large number of accidents. There is thus a one-to-many relationship between tables *magDetail* and *accDetail*.

Even if Tables A and B do not share a common attribute they can still be related. If Table A is related to Table C and Table C is related to Table B then there is a relationship between Tables A and B. This can be illustrated by the addition of a table called *regDetail* (see Fig. 2.5). The *accDetail* and the *magDetail* have the *magArea* attribute in common while the *magDetail* and the *regDetail* have the *Region* in common. If the magisterial area in which an accident occurred is known it is possible to determine the region from *magDetail*, and once the

Region	Province
S Cape	W Cape
Boland	W Cape
W Coast	W Cape
Metro	W Cape

Figure 2.5 : A database table

region is known it is possible to determine in which province the accident happened from the *regDetail* table. There is an indirect one-to-many relationship between *accDetail* and *regDetail*.

c) Many-to-many relationship

If a record in Table A is related to zero or more than one records in Table B and vice versa there exist a many-to-many relationship between Tables A and B.

2.5.2 DATA EXTRACTION and REDUCTION

Having good quality accident data in a well-designed and fully functional database is worthless unless the data can be accessed, extracted and reduced in an efficient manner for the purposes of reporting and analysis.

The most common method to extract information from a database is using SQL (Structured Query Language).

The objectives of this section are :

- To introduce the reader to SQL.
- To show how SQL can be used to extract information.
- To show how SQL can be used to do basic analysis.

SQL is a 4th generation non-procedural language. It provides tools for a variety of data management tasks :-

- Querying of data
- Inserting, updating and deleting records

Data is queried by formulating and executing, within your software system, a SQL procedure which in its simplest form has the following general format.

Select [column names/ * /expression]
From [table names]
Where [join condition/ criteria/expression]
And/Or [criteria/expression]
Group by [column names]
Order by [column names] **desc/asc**

To illustrate the use of SQL procedures to extract and reduce data the Tables in Figure 2.6 will be used. Table *AccDetails* contains details on individual accident records. Table *LinkDetails* contain details on geometric and traffic flow characteristics of link segments. *AccDetails* and *LinkDetails* have the *Roadno* attribute in common. There exist a many-to-one relationship between *LinkDetails* and *AccDetails* i.e. a record in *AccDetails* can only be related to one record in *Linkdetails* while one record in *Linkdetails* can be related to many records in *AccDetail*. Since a road with a specific *Roadno* may consist of more than one link the two tables can only be joined by also specifying in the joining operation that a record in *AccDetails* is related to that segment in *LinkDetails* for which *Acckm* falls between *Startkm* and *Endkm*.

For the purposes of this exercise is assumed that the *Climate* (e.g dry, medium, wet) depends on the magisterial area in which the link fall. Each magisterial area is associated with one climate type and one only. The *LinkDetails* and *AreaDetails* tables have the *magArea* (Magisterial area) in common.

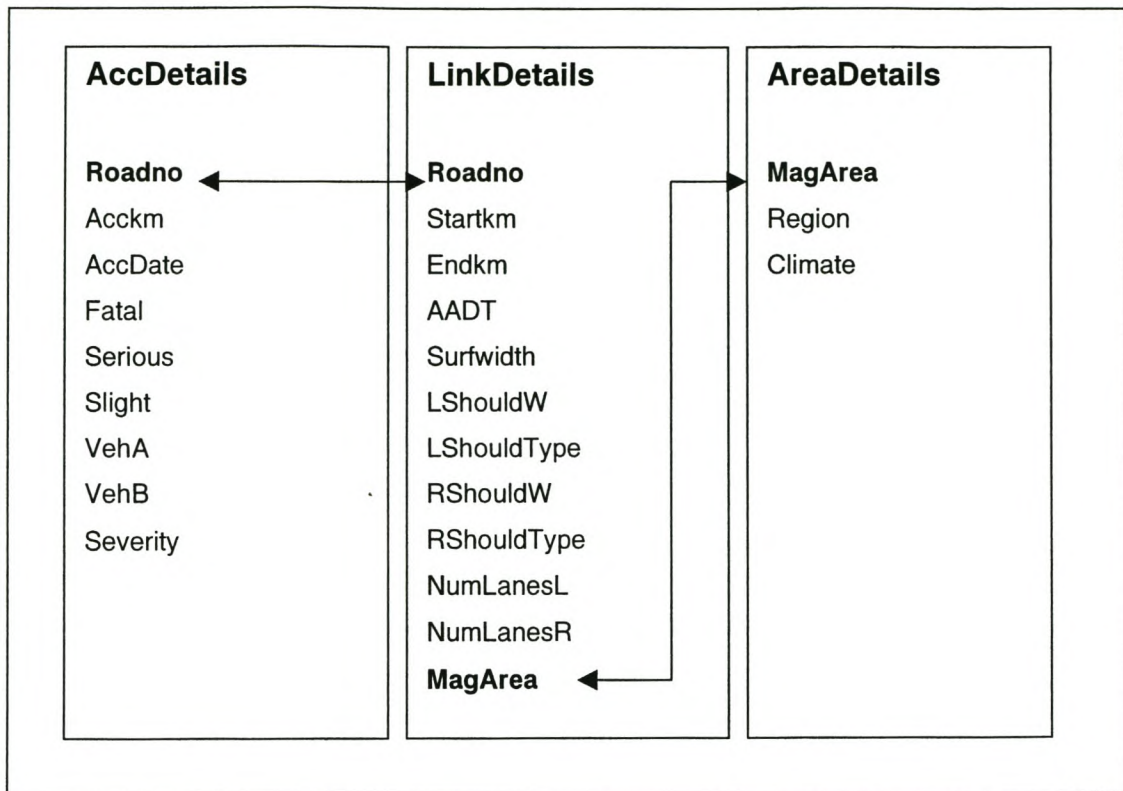


Figure 2.6 : Normalised database tables

a) SELECT statement

A SELECT statement retrieves information from a database. In its simplest form it must include

- A SELECT clause which specifies the columns to be retrieved and/or the expression to be executed
- A FROM clause which specifies the table/s on which these columns exist.

Ex 2.2: To retrieve all attributes of all records in the *AccDetails* table.

```
SELECT *
FROM AccDetails;
```

Ex 2.3 : To return the road number and accident date attributes of each record in the *AccDetails* table.

```
SELECT RoadNo, AccDate  
FROM AccDetails;
```

Ex 2.4 : To return the road number, date and total number of casualties i.e. the sum of the fatal, serious and slight attributes for all accidents in the *AccDetails* table.

```
SELECT RoadNo, AccDate, (Fatal+Serious+Slight)  
FROM AccDetails;
```

b) WHERE statement

The WHERE clause is optional and is used to i) join two or more tables
ii) or to specify the selection of certain rows.

i) Joining two or more tables

Ex 2.5 : To return the *AADT* and the *surfaced width* at the location of all accidents, the *AccDetails* and *LinkDetails* can be joined as follows :

```
SELECT a.*, l.AADT, l.SurfWidth  
FROM accDetails a, LinkDetails l  
WHERE a.Roadno = l.Roadno;  
AND a.Acckm between l.Startkm and l.Endkm;
```

Ex 2.6 : To return the *AADT*, the *surfaced width* and the *climate* for all accidents, the *AccDetails*, *LinkDetails* and *AreaDetails* can be joined as follows:

```
SELECT a.*, l.AADT, l.SurfWidth, r.Climate  
FROM accDetails a, LinkDetails l, AreaDetails r  
WHERE a.Roadno = l.Roadno  
AND l.Area = r.Area  
AND a.Acckm between l.Startkm and l.Endkm;
```


In the above SQL procedures the letters *a*, *l* and *r* are referred to as aliases.

ii) Specifying the selection of rows

The WHERE statement can be used to limit the display of records by specifying certain conditions. The WHERE clause consists of three elements

- a) A column name
- b) A comparison operator (<, >, >=, <= or =)
- c) A column name, constant or list of values

Ex 2.8 : To return all accidents in the *AccDetails* table where there are 5 or more fatalities.

```
SELECT a.*
FROM AccDetails a
WHERE a.Fatal >= 5;
```

Ex 2.9 : To return all accidents in the *AccDetails* table where there are more fatalities than injuries.

```
SELECT a.*
FROM AccDetails a
WHERE a.fatal > a.serious + a.slight;
```

Ex 2.10 : To retrieve all accidents in the *AccDetails* table that occurred on roads with a surfaced width of greater than 10m and an AADT greater than 3500 vehicles per day :

```
SELECT a.*
FROM AccDetails a, LinkDetails l
WHERE l.Roadno = a.Roadno
AND a.Acckm between l.Startkm and l.Endkm
AND l.AADT >= 3000
AND l.SurfWidth >= 10;
```

c) GROUP BY Statement

The GROUP BY statement allows for summary information to be obtained for groups of rows through the use of aggregate functions. It divides rows into smaller groups.

The following group functions can be used in the SELECT statement when using the GROUP BY statement :

AVG[column/expression] - Return the average value of *column/expression* over all specified records.

COUNT[]* - To count all records per group that meet the specified criteria.

MAX/MIN[column/expression] - Return the maximum/minimum value of a *column/expression* per group for all records that meet the specified criteria.

SUM[column/expression] - Return the sum of a *column/expression* per group for all records that meet the specified criteria.

The *STDDEV* and *VARIANCE* functions are similar to the SUM function and return the standard deviation and variance of a *column/expression* respectively.

Ex 2.11 : To determine the total number of accidents per link.

```
SELECT I.Roadno, I.Startkm, I.Endkm, Count(*)
FROM AccDetails a, LinkDetails I
WHERE I.Roadno = a.Roadno
AND a.Acckm between I.Startkm and I.Endkm
GROUP BY I.Roadno, I.Startkm, I.Endkm;
```

Ex 2.12 : To determine the number of fatal, serious, slight and damage only accidents for each link.

```
SELECT I.Roadno, I.Startkm, I.Endkm, a.Severity, Count(*)
FROM AccDetails a, LinkDetails I
WHERE I.Roadno = a.Roadno
AND a.Acckm between I.Startkm and I.Endkm
GROUP BY I.Roadno, I.Startkm, I.Endkm, a.Severity;
```

Ex 2.13 : To determine the number of casualties per link that resulted from accidents involving at least one heavy vehicle.

```
SELECT I.Roadno, I.Startkm, I.Endkm, Sum(Fatal+Serious+Slight)
FROM AccDetails a, LinkDetails I
WHERE I.Roadno = a.Roadno
AND a.Acckm between I.Startkm and I.Endkm
AND a.VehA like 'Heavy_Vehicle' OR a.VehB like 'Heavy_Vehicle'
GROUP BY I.Roadno, I.Startkm, I.Endkm;
```

Ex 2.14 : To calculate the accident rate for each link.

```
SELECT I.Roadno, I.Startkm, I.Endkm, [Endkm-Startkm]* AADT*3*
      365/106 as E, Count(*) as A, (A)/(E) as R
FROM AccDetails a, LinkDetails I
WHERE a.Roadno and I.Roadno
AND a.Acckm between I.Startkm and a.Endkm
AND a.Accdate between '01-JAN-93' and '31-DEC-95'
GROUP BY I.Roadno, I.Startkm, I.Endkm, [Endkm-Startkm]*AADT*3*
      365/106;
```

Ex 2.15 : To calculate the total exposure for links with surfaced shoulders and a surfaced width greater than 11 m (assuming a 3 year period).

```
SELECT sum[(I.Endkm-I.Startkm] *3*365* I.AADT/106)] as E
FROM LinkDetail I
WHERE I.Surfwidth >= 11
AND I.LShouldType like 'Surfaced'
AND I.RShouldType like 'Surfaced'
```

Ex 2.16 : To calculate the total number of accidents on links with surfaced shoulders and a surfaced width of >= 11m.

```
SELECT count(*)
FROM AccDetail a, LinkDetail I
WHERE I.Roadno = a.Roadno
AND I.LShouldST like 'Surfaced'
AND I.RShouldST like 'Surfaced'
```

2.6 SUMMARY and CONCLUSION

This Chapter identified three major information sources required to successfully implement a road safety remedial programme ; accident data, network and environmental data and, traffic flow information.

Various issues that could affect the quality of accident data in South Africa were identified. The legislated definition of an accident was presented as well as other supporting pieces of legislation to show that the legal definition could be open to misinterpretation which could cause underreporting of accidents. The issue of under reporting was identified as a major problem that could seriously impact on the efficiency of a road safety remedial programme. The results of international research studies on the extent of under reporting relating to different road user groups and severity classes were presented. A methodology to quantify the effect of under reporting on the accuracy of safety estimates was presented.

Reference were made to research studies that found that the safety at a location depends, amongst others, on its geometric and environmental characteristics. Many accident data methodologies require the use of a reference group or a comparison group, which necessitate the availability of road network and environmental characteristics. Without this information it would not be possible to partition road network elements into reference groups.

Traffic flow information was identified as important because this type of information is required to estimate levels of exposure and to develop multivariate regression models such as Safety Performance Functions. Issues regarding the accuracy of AADT estimates were discussed . It was reported that the accuracy of an AADT estimate depends on the stratification method, day/s of the week counted and the length of the sampling period. The different stratification methods were discussed and coefficient-of-variation values for each were presented for the different sampling periods and day/s of the week.

An overview was provided on ways to manage these information sources in a manner that would ensure that data required for a safety study is readily available in the desired format. A method to extract data from a number of relational databases containing accident, network and count data was presented. A number of illustrated examples were given to show how SQL (Structured Query Language) can be used to extract data, suitable for analysis, from a number of databases.

CHAPTER 3

MEASURING SAFETY

3.1 INTRODUCTION

The ability to obtain reliable and valid estimates of safety for a variety of entity types is fundamentally important for the efficient identification and ranking of hazardous locations and to successfully determine the effect of various road safety remedial measures.

The primary objective of this chapter is to investigate and to report on various issues to consider before subjecting accident data to a formal examination in order to estimate the safety of various entity types.

Firstly, the issue of what 'safety' is and how it can be defined will be addressed. In order to facilitate an improved understanding of what is meant by 'safety' the concepts of *reliability*, *validity* and *measurement error* will be discussed. Certain recommendations will be made on how to reduce *measurement error* and to increase the *validity* and *reliability* of a safety estimate.

One of the most important determinants of the validity of a safety estimate – the *accident measure* i.e. what type and kind of accidents to use, will be discussed in detail.

The validity and reliability of a safety estimate expressed as an accident rate, in addition to the accident measure, also depends on the measure of *exposure* used to calculate the accident rate. For an accident rate to be a valid and reliable measure of safety it should meet certain requirements. These requirements will be presented as well as the potential consequences of violating these requirements. It will be shown that conventional accident rate expressions commonly used for segments and intersections do not meet

the requirements and are therefore not suitable to estimate the safety of road segments and intersections.

To overcome the shortcomings of using conventional accident rates an alternative approach to estimating safety will be presented. This approach, it will be shown, relies heavily on Safety Performance Functions (SPF) i.e. mathematical functions that relate traffic flows to accidents.

Next, an extensive overview of typical Safety Performance Functions for segments and intersections will be presented and will be illustrated with case studies. It will be shown that using Safety Performance Functions to estimate safety, in certain circumstances, also has certain shortcomings which could impact on the validity and reliability of safety estimates. These shortcomings - the Function Averaging Problem and the Argument Averaging Problem will be discussed in detail.

3.2 DEFINING SAFETY

According to Hauer (1997) the principle manifestation of safety (or 'unsafety') are accidents and the harm they cause i.e. damage to property, loss of life, injury etc.

Accident data is therefore the primary source of information to make inferences about road safety.

The analysis of accident data in the planning and evaluation of a road safety accident remedial programme requires first, the comparison of safety levels to identify and rank hazardous locations, and secondly to compare the level of safety between the 'before' and 'after' periods of an entity or a group of entities.

The quality of the results of such analysis and therefore the overall integrity of a road safety remedial programme depends directly on the *reliability* and *validity* of the safety estimates used in the analysis.

Hauer (1997) provides the following definition of safety at an entity :

*“The number of accidents, or accident consequences, by kind and severity, **expected** to occur on the entity during a specified period of time.”*

Often in practice safety is measured not by expected accident frequency but by the **expected accident rate**.

According to Mahalel (1986) a common method of defining the safety of an entity is by means of risk and exposure. Risk is defined as the possibility of experiencing a ‘negative’ event, such as, for example, a road traffic accident. In statistical terms the risk associated with a transportation system represents the probability of being involved in an accident while using the system. Exposure is defined as the amount of risk (expressed as the total number of opportunities for an event to occur) a road user exposes him or herself to. The total expected number of accidents to occur at a system during a certain time period is given by the product of the risk and exposure.

$$A = R * E \qquad \dots[3.1]$$

A - Expected number of accidents in time period T.

R - Risk : The probability of an accident.

E - Exposure : A measure of the total number of opportunities for an accident to occur during time period T.

The term ‘**expected**’ is used here to refer to an estimate that is both reliable and valid and hence free from *measurement error* .

3.3 MEASUREMENT THEORY

Leedy (1993) provides the definition of what *measurement* is :

“Measurement is limiting the data of any phenomenon – substantial or – insubstantial – so that data may be examined mathematically, and ultimately, according to an acceptable qualitative or quantitative standard.”

In order to research any road safety problem, of whatever nature, the first step is to identify indicator/s that will provide an acceptably reliable and valid quantitative measure of the research problem. Such indicator/s form the basis of the research process. (Leedy ; 1993)

3.3.1 VALIDITY AND RELIABILITY

According to Leedy (1993) with any type of measurement *validity* and *reliability* are two very important considerations.

Validity is concerned with the soundness and the effectiveness of the safety estimate. It raises the question whether an estimate measures what it is suppose to measure, as well as how comprehensively and accurately. An important determinant of the validity of a safety estimate is therefore the chosen ‘accident measure’ i.e. the type of accidents to consider in the estimation of safety. Issues relating to ‘accident measures’ are covered in more detail in Section 3.4.

Reliability deals with how dependable and accurate an estimate is. A major determinant of the degree of reliability is the extent to which the estimate contains any *random* and *systematic* errors (see Section 3.3.2). The larger the degree of these errors the less the reliability of the estimate.

An indicator needs to be reliable for it to be valid. (Leedy ; 1993)

3.3.2 MEASUREMENT ERROR

Any measurement (X) or observed value consists of 3 different components :

- a) True perfect value (T)
- b) Systematic Error (S)
- c) Random Error (R)

Where $X = T + S + R$

Under conditions of perfect validity there are no systematic and random errors i.e. $X = T$. The validity of a measurement can be improved by focussing on eliminating S and R.

i) RANDOM ERROR

Accidents are discrete random events that according to Abbes et al. (1981) follow a Poisson distribution around the 'true mean' (T). It is unlikely therefore that the observed number of accidents X at an entity will be equal to the true mean (T). Assuming the systematic error = 0, the difference between T and X represents the random error associated with a measurement.

ii) SYSTEMATIC ERROR

Whereas random error is a statistical phenomenon, systematic error is the result of external influences, which distorts a measurement in a systematic way.

The most common causes of systematic errors relating to accident data are :

- a) The underreporting of accidents.
- b) Inaccurate and incorrect accident information.
- c) Missing accident information.

The following remedial measures can be taken to minimise *systematic errors* :

- a) Redefine the chosen 'accident measure' so as to exclude those subsets of data that are normally associated with significant systematic errors - such as *Damage Only* accidents.
- b) Systematic errors as a result of inaccurate or missing data can really only be reduced by having a good accident data management system in place. This issue is covered in more detail in Chapter 2.

3.4 ACCIDENT MEASURE

An *accident measure* refers primarily to the type/s of accidents to consider when estimating safety. It refers to what is called primary data – data that lies closest to the source of 'ultimate truth' – data that will reveal most about the true nature of the research objective. It refers to that subset of available data that will have maximum validity and reliability when compared to other subsets.

When deciding on an appropriate accident measure it is important to consider the objective of the analysis. As far as road safety accident remedial programmes are concerned the objectives could be to identify hazardous locations or to evaluate the effectiveness of road safety remedial measures.

When evaluating road safety treatments it is necessary to identify the kind of accidents that will be affected by the treatment. Hauer (1997) refers to these accidents as *target accidents*.

Hauer (1997) provides the following definition for *target accidents* :

“The target accidents of a treatment are those accident types the occurrence of which can be materially affected by the treatment.”

Comparison accidents, in contrast, is defined by Hauer (1997) as follows :

“Comparison accidents for a treatment are those accidents the occurrence of which cannot be materially affected by the treatment.”

Identifying *target accidents* require an understanding of the accident generation process i.e. how the treatment works.

Accidents at an entity can be grouped into different subsets each of which are associated with a different level of risk. For example research by Persaud and Musci (1995) has indicate that for rural 2 lane segments in Ontario, Canada, there are at least 4 distinct accident data types, i) night-time single vehicle accidents, ii) night-time multi-vehicle accidents, iii) day-time single vehicle accidents and iv) day-time multi-vehicle accidents. Each of these accident types is associated with significantly different levels of accident risk, with the risk of night-time accidents being considerably higher than that of day-time accidents.

The magnitude of an aggregated accident measure which include all day-time and night-time accidents, will depend on amongst others on the ratio of night-time to day-time traffic flows. Comparing different entities, each of which might have a different night-time / day-time traffic ratio, becomes problematic. One entity, for example, might be identified as more hazardous than another only because it carries relatively more night-time traffic and not because it is inherently more unsafe.

There are therefore advantages in using, instead of a single ‘aggregated’ ‘accident measure’, a number of ‘accident measures’ to estimate the safety at a location.

The guiding principle is that ideally all accidents included in the ‘accident measure’ should be associated with the same level of risk.

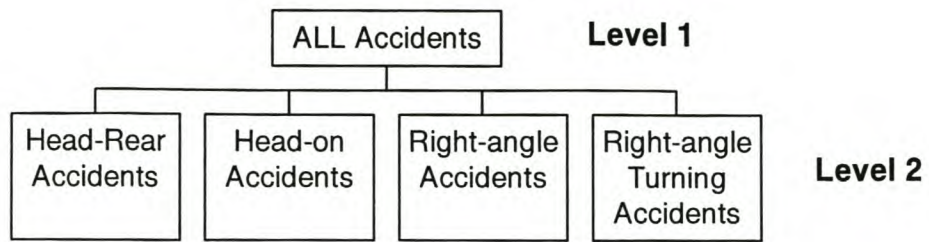


Figure 3.1 : Typical categorisation of intersection accidents.

Figure 3.1 indicates two levels of data aggregation at an intersection where Level 1 = ALL accidents and Level 2 = 4 different accident types. Intersections are often dominated by a specific accident type. Using a Level 1 aggregation this will not be apparent. The different accident types as identified in aggregation Level 2 do not necessarily share the same level of risk and the combination thereof into a single level of safety for an intersection can produce a result that is not valid nor reliable due to the bias introduced by the *Function Averaging Problem*. (See Section 3.7.4)

An important consideration when choosing an ‘accident measure’ is that different levels of data aggregation have different information requirements, especially for developing appropriate *exposure* measures. For example, in estimating the safety of an intersection using ALL accidents as an accident measure only requires information on traffic flows entering the intersection. However turning movement traffic volumes are required should the objective be to use Level 2 accident measures to define the safety of an intersection.

At lower levels of aggregation the sub-sets become more homogenous but this is achieved at the expense of sample size. Because of the of the high variability inherent in accident data large sample sizes are generally required to obtain reliable safety estimates. There should be a proper balance between the demands for homogeneity and sample size requirements.

3.5 EXPOSURE

The most widely used means of describing risk is the accident rate. Accident rates are defined as the number of accidents (as defined by the 'accident measure') divided by an exposure measure.

The use of accident rates is based on the assumption that there is always a perfectly linear relationship between A and E . In other words, the accident rate is completely independent of the level of exposure (E).

The inclusion of an exposure measure in estimating accident risk is required to equalise for differences in the intensity of use in order to make comparisons more meaningful. Accident rates determined in this manner are used to standardise safety with respect to traffic flow. According to Chang (1982) unless the exposure is known the relative hazards of various situations cannot be compared.

A valid and reliable measure of exposure should meet the following criteria :

- a) It should be a direct measure of the total number of opportunities for an accident (as defined by the accident measure) to occur during the study period.
- b) If a 'proxy' measure is used because it is not possible to determine the total number of opportunities with precision than this 'proxy' measure should be directly proportional to the total number of accident opportunities.
- c) For it to be a direct measure the exposure measure should be linearly related to the accident measure. According to Hauer (1995) in order for exposure to do the job of equalising the Safety Performance Function (SPF)(i.e. the relationship between the two measures) must be linear.

According to Hauer (1995) choosing a measure of exposure that is not linearly related to the accident measure could have the following consequences :

- a) When the SPF is not a straight line, the accident rate will change as the amount of traffic (exposure) changes, even if there was no intervention and the road remained the same. It is possible for the accident rate to decrease even as the facility becomes less safe. It is even possible when two facilities are compared with each other for the safer facility to have a higher accident rate than the other facility.
- b) When evaluating the effectiveness of a remedial measure a non-linear SPF could cause the effectiveness to be over or under estimated. This is illustrated in Figure 3.2.

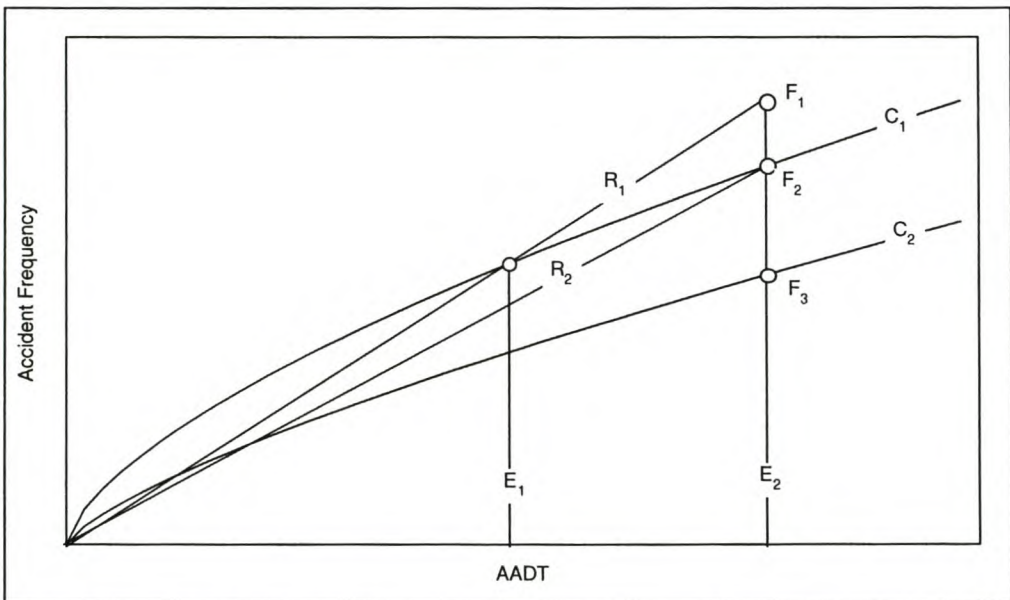


Figure 3.2 : Non-linear Safety Performance Function

The accident rate is obtained by the slope of the line connecting the origin with any particular point on the curve. At exposure level E_1 the accident rate is R_1 . If the exposure level was to increase to E_2 the accident rate will be R_2 . It is evident that $R_2 < R_1$. In other words in the absence of any road safety improvements an increase in the exposure level in this case led to a reduction in the accident rate.

In the event of successfully improving the safety of the road the SPF will shift downwards – from Curve 1 to Curve 2. The improvements could possibly attract more traffic which would cause the exposure to increase from E_1 to E_2 . The evaluation of a safety measure requires an estimate of what the number of accidents would have been had no improvements been undertaken. Assuming a linear SPF (described by R_1) this estimate would be given by point F_1 where in fact the correct estimate would be given by point F_2 . The real safety effect would therefore be given by the difference between F_2 and F_3 and not by the difference between F_1 and F_3 . Assuming a linear SPF would cause the safety effect to be overestimated by the difference between F_1 and F_2 .

3.5.1 ROAD SEGMENTS

The most widely used measure of exposure for measuring safety on road segments is the million-vehicle-kilometres measure that is determined as follows :

$$E = AADT * L * n * 10^{-6} \quad \dots[3.2]$$

E - Exposure in million-vehicle-kilometers (mvkm)

$AADT$ - Average Annual Daily Traffic (veh/day)

L - Length of segment in kilometres (km)

n - Number of days of study period.

There is consensus (Hauer ; 1995, Mahalel ; 1986 and Satterwaith ; 1981) that using this measure of exposure will result in an accident rate that is not reliable for the following reasons :

- a) The relationship between E and most *accident measures* (e.g. all accidents, single vehicle accidents, multi-vehicle accidents etc.) was found to be non-linear.

Satterwaithe (1981) in a survey of research into relationships between traffic accidents and traffic volumes concluded that the weight of evidence suggests that single vehicle accidents and multi-vehicle accident rates depend in a fundamentally different way on traffic volumes. The single-vehicle rate tends to decrease with increasing traffic volumes while on the other hand the multi-vehicle rate increases with increasing traffic volumes, with some evidence of a decrease in the rate after a certain volume, perhaps corresponding to the onset of congestion. The implications of these relationships is that the total accident rate varies in a U-shaped fashion with the traffic volume, but the form of this relationship is likely to vary substantially, depending on the relative numbers of single and multi-vehicle accidents. Satterwaithe (1981) however does acknowledge that there are a lot of conflicting results concerning the relationship between the multi-vehicle rates and traffic volumes.

Zhou and Sissipiku (1997) developed the following models to relate accident rates and exposure (expressed as a volume-to-capacity ratio) on urban freeways in Michigan. (See Figure 3.3)

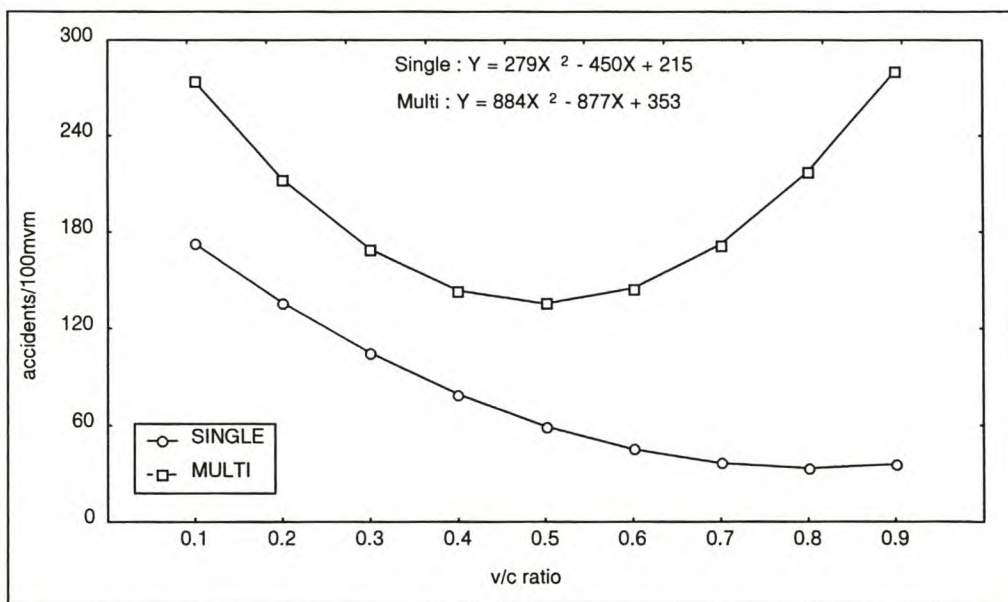


Figure 3.3 : SPF's for single and multi-vehicle accidents on Michigan freeways

Zhou and Sissipiku (1997) explains the form of these relationships in the following manner :- . As traffic volumes increase so do vehicle conflicts and therefore higher accident rates become impossible. To explain the higher accident rates associated with low traffic flows one has to consider that low traffic flows generally occur at night where poor visibility, fatigue and higher rates intoxication are major factors. Also in low exposure conditions the average driver 's attention to driving tasks is reduced and higher speeds may be selected.

In conclusion, there is therefore sufficient empirical and logical support to the notion that the relationship between A and E (as in Equation 3.2) to be non-linear.

- b) In the exposure function a time-average value of traffic flow, the AADT is used. There exist a causal relationship between traffic flow and accidents. The effect (accidents) is observed over a long period of time during which the cause (traffic flows) has assumed widely different values, but for which only the average value is known. The causal link between accidents and AADT is therefore indirect. This leads to what is referred to as the issue of *argument averaging* which could introduce, depending on the form of the SPF, a large and significant bias. (Mensah and Hauer ; 1998)

3.5.2 INTERSECTIONS

A common measure of exposure to measure safety at intersections is the *sum-of-flows* measure that is determined as follows :

$$E = d * \sum AADT_{in} * 10^{-6} \quad \dots[3.3]$$

$\sum AADT_{in}$ - *Sum of incoming flows on all approaches (veh/day).*

d - *Number of days of study period.*

The *sum-of-flows* measure for intersections determines the total number of vehicles that entered the intersection during the study period. This does not and cannot account for possible correlation between specific accident types and certain combinations of vehicle movements. It also implies that all vehicles entering an intersection have an equal probability of being involved in an accident (Plass et al. ; 1987).

According to Hauer et al. (1988) the assumption that the number of accidents at an intersection is proportional to the sum of flows that enter the intersection is logically unsatisfactory and not a suitable basis for engineering analysis which attempts to link cause and effect. Hauer et al. (1988) suggest that one could expect, for example, that the number of rear-end accidents at an intersection approach will strongly depend on the flow on approach A and depend only weakly on the flow on approaches B, C and D. One would also expect that accidents between vehicles from traffic streams moving at right-angles will be related to the product of these flows. To use the sums of these flows leads to the logical difficulty that one will be able to predict accidents even if one of the flows is zero.

3.6 AN ALTERNATIVE APPROACH TO SAFETY ESTIMATION

On the basis of the fact that the conventional concept of accident risk, as determined by the accident rate, is a function of the exposure level – the use of the accident rate, in some cases, could be an inappropriate measure of risk. According to Mahalel (1986) the risk level can only be expressed in relation to a specific exposure level. Using accident rates without relating them to exposure levels could make the comparison of accident rates for the purposes of identifying and ranking hazardous locations and evaluating remedial measures meaningless.

According to Hauer (1995) two accident rates can only be compared with each other if there exists a reasonable expectation that the two rates should

be equal to each other. This expectation is not necessarily valid when two accident rates were determined using different levels of exposure.

The problems associated with the conventional definition of accident risk can be overcome by relating it to exposure. This can be achieved by using Safety Performance Functions.

Mahalel (1986) proposed that the probability of a certain number of accidents at a given level of exposure to be an alternative definition of accident risk. In terms of Mahalel's alternative definition of risk, the risk (R) at any location is a function of the exposure (E), the expected number of accidents (A) and the probability of A accidents (P_a).

$$R = f(E, A, P_a) \quad \dots[3.4]$$

The fundamental characteristic of this alternative definition of accident risk is its ability to express the expected number of accidents or the probability of a certain number of accidents at any exposure level. Accordingly, the risk level of a system can only be expressed in relation to a specific exposure level. (Mahalel ; 1986)

The role of a SPF can be seen as a 'black box' where exposure is an 'input' and in which 'output' is accidents and probabilities.

3.7 SAFETY PERFORMANCE FUNCTIONS

Safety Performance Functions are in essence multivariate regression models that explain how the number of accidents at a location depend on the traffic flows.

3.7.1 ROAD SEGMENTS

Mensah and Hauer (1988) mentions four common forms of SPF's commonly used to relate accidents to exposure on roadway segments :

a) Type 1 : The Exponential SPF :

$$A = \alpha (ADT)^\beta \quad \dots[3.5]$$

This function assumes that the exposure (total number of opportunities) is proportional to ADT^β . The value of A will always increase with increasing values of ADT.

The accident rate (R) at any level of ADT is given by :

$$R = \alpha (ADT)^{\beta-1} \quad \dots[3.6]$$

If $\beta > 1$ then R will always increase with increasing values of ADT, and if $\beta < 1$ then R will always decrease with increasing values of ADT. Please refer to Figure 3.4.

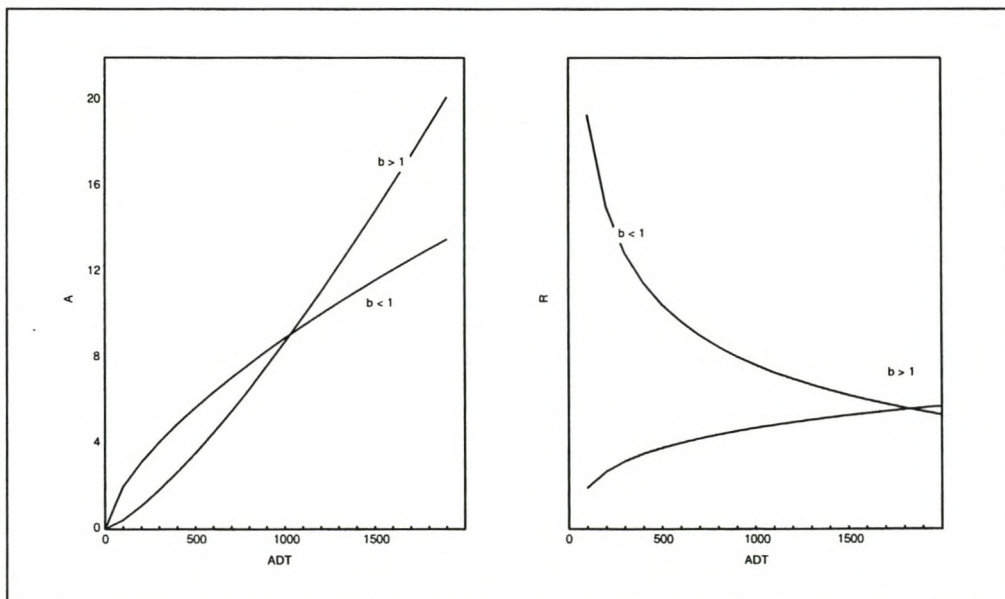


Figure 3.4 : Accident number and rate curves : Exponential SPF

The exponential SPF has been used by Persaud and Musci (1995) to model the relationship between ADT and accidents on two-lane rural roads in Ontario, Canada.

Persaud and Musci (1995) used data from Ontario, Canada, for two-lane rural road segments to develop a series of SPF's using the Exponential form (Type1).

$$A = \alpha Q^\beta \quad \dots[3.7]$$

- A - *accidents/kilometre/year*
- Q - *Average hourly traffic flow*

Table 3.1 : Regression parameters of SPF's

Accident Type	Day-time		Night-time		24 Hours	
	α	β	α	β	α	β
Single-vehicle	0.0040	0.490	0.0354	0.557	0.0657	0.444
Multi-vehicle	0.0008	1.173	0.0013	1.071	0.0011	1.123
All	0.0194	0.741	0.0468	0.650	0.0415	0.627

The study by Persaud and Musci (1995) has shown that on 2-lane roads in Ontario, Canada there exist significantly different SPF's for day-time single vehicle accidents, day-time multi- vehicle accidents, night-time single vehicle accidents and night-time multi-vehicle accidents.

The Safety Performance Functions developed by Persaud and Musci (1995) are illustrated in Figures 3.5 and 3.6.

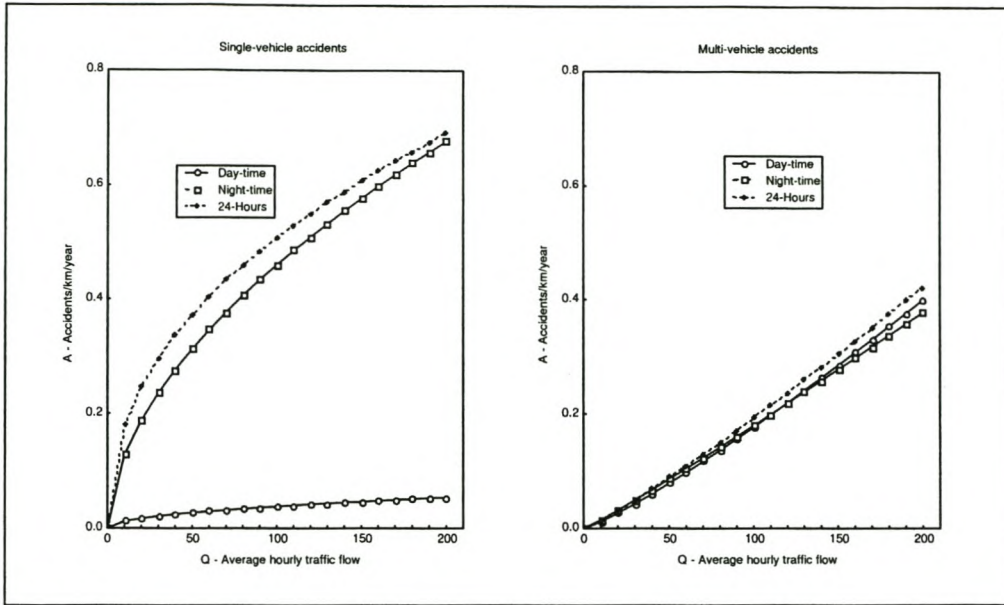


Figure 3.5 : Single vehicle and multi-vehicle accident SPF's

From Figure 3.5 it is evident that at similar levels of exposure the risks associated with single vehicle accidents are considerably worse during the night than during the day.

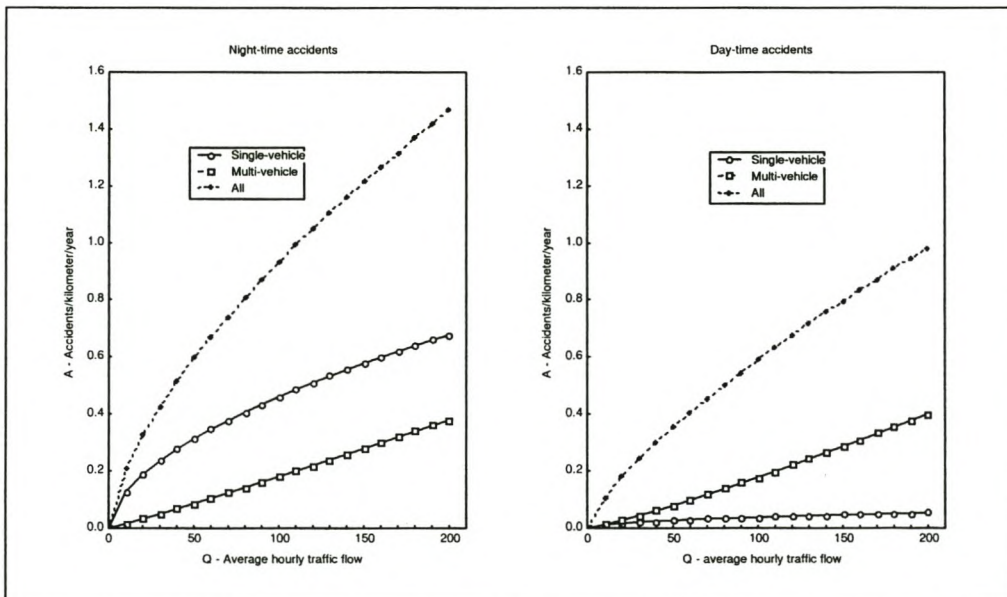


Figure 3.6 : Night-time and day-time accident SPF's

From Figure 3.6 it is evident in night-time conditions the risk associated with single vehicle accidents are considerably worse than that of multi-vehicle accidents while the reverse applies in day-time conditions.

Persaud (TRR1327) in a study to estimate the accident potential of Ontario road sections developed the following SPF's for different classes of roads.

$$A = \alpha Q^\beta \quad \dots[3.8]$$

A - Accidents/kilometre/year

Q - ADT (1000 vehicles)

Table 3.2 : Regression parameters of SPF's

Road class	α	β	k
Freeway	0.6278	1.024	2.95
Rural/undivided/2-lane	1.3392	0.8310	2.90
Rural/undivided/multi-lane	0.6528	1.3037	2.90
Urban/undivided/2-lane	3.6514	0.5588	2.90
Urban/undivided/multi-lane	1.4196	0.8763	2.90
Rural/divided/multi-lane	0.4591	1.3037	2.90
Urban/divided/multi-lane	0.9984	0.8763	2.90

k – Dispersion parameter obtained from Negative Binomial regression.

The results of Persaud's study have shown that different types of facilities in Ontario have significantly different SPF's.

b) Type 2 : The Quadratic SPF

$$A = \alpha (ADT) + \beta (ADT)^2 \quad \dots[3.9]$$

The accident rate at any value of ADT is given by :

$$R = \alpha + \beta (ADT) \quad \dots[3.10]$$

Figure 3.7 shows the accident number and accident rate curves for the following quadratic SPF : $A = 0.0005*Q - 2*10^{-8}*Q^2$.

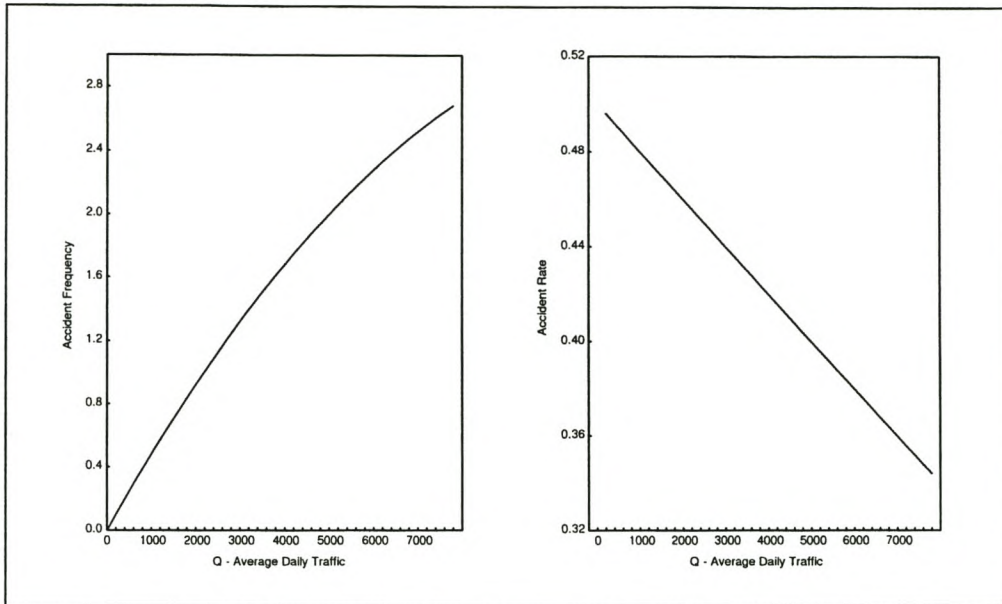


Figure 3.7 : Accident number and rate curves : Quadratic SPF

The Safety Performance Function developed by Zhou and Sissipiku (1997) as shown in Figure 3.3 is of the Quadratic type.

c) Type 3 : Hoerl’s Function with k = 1

According to Mensah and Hauer (1998) Hoerl’s function is as follows :

$$A = \alpha (ADT)^k e^{\beta (ADT)} \quad \dots[3.10b]$$

Thus with k = 1, then :

$$A = \alpha (ADT) e^{\beta (ADT)} \quad \dots[3.11]$$

The accident rate at any value of ADT is given by :

$$R = \alpha e^{\beta (ADT)} \quad \dots[3.12]$$

With small values of ADT the accident frequency (A) increases approximately linearly with traffic flow. Mensah and Hauer (1998) noted that this function may be suitable for single vehicle accidents.

d) Type 4 : Hoerl's function with k = 2

$$A = \alpha (ADT)^2 e^{\beta(ADT)} \quad \dots[3.13]$$

The accident rate at any value of ADT is given by :

$$R = \alpha (ADT) e^{\beta(ADT)} \quad \dots[3.14]$$

A increases with the square of traffic flow. At large traffic flows the slope of the function begins to diminish. Eventually a peak is reached and A will begin to decrease. Mensah and Hauer (1988) noted that this function may be suitable for two-vehicle accidents.

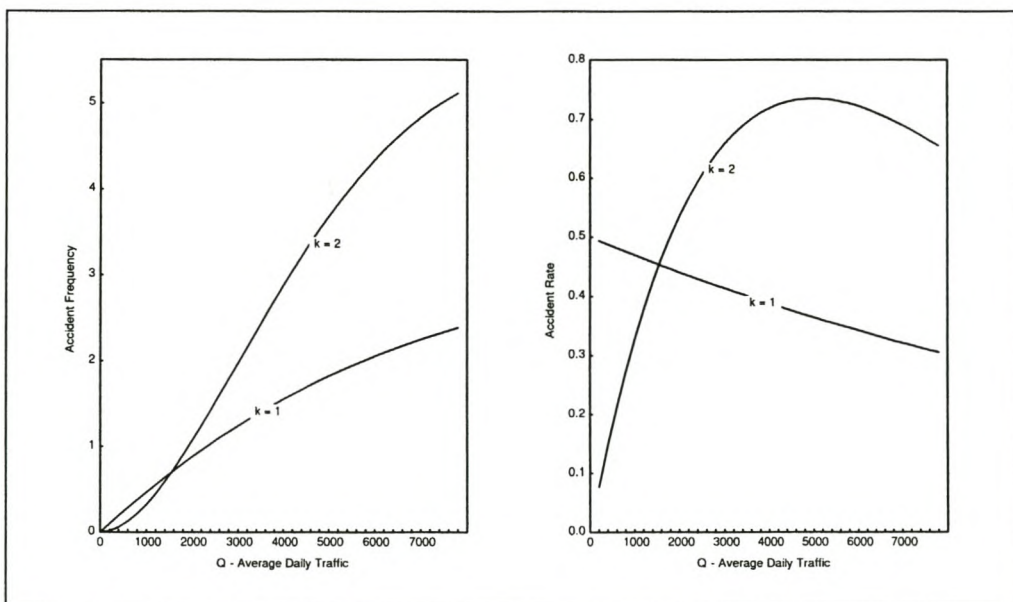


Figure 3.8 : Accident Number and Rate curves : Hoerl's function with k =1 and 2.

3.7.2 INTERSECTIONS

Some researchers such as Breunning and Bone, Sunti and Hakkert, and Mahalel (in Hauer et al. ; 1988) related accidents to the products of the conflicting flows.

$$A = \alpha Q_1 Q_2 \quad \dots[3.15]$$

Empirical research according to Hauer et al. (1988), has found the above relationship not to be correct. It was found that accidents are rather related to the product of flows with each flow raised to a power of less than 1. Tanner (in Hauer ; 1988) suggested that the square root of the product flows would be sufficiently accurate as a rule of thumb :

$$A = \alpha \sqrt{Q_1 Q_2} \quad \dots[3.16]$$

The '*products-of-flows-to-power*' relationship has been used by Bonneson and McCoy (1993) and Belanger (1994) to relate total accidents to total flows on the major and minor approaches at two-way stop controlled intersections on rural highways, while Hauer et al. (1988), in a study to estimate safety at signalised intersections used the '*products-of-flows-to-power*' relation to relate the frequency of specific accident types to the relevant conflicting flows.

Bonneson and McCoy (1993) developed the following SPF for two-way stop controlled intersections on rural highways in Minnesota (USA) :

$$A = 0.692 \left(\frac{T_m}{1000} \right)^{0.256} \left(\frac{T_c}{1000} \right)^{0.831} \quad \dots[3.17]$$

T_m - major road traffic flow (veh/day)

T_c - Minor (cross) road traffic flow (veh/day)

A - Annual expected accident frequency (within 153 m / 500ft of junction)

k - Dispersion parameter from Negative Binomial regression

For 4-legged un-signalised intersections in eastern Quebec, Bélanger (1994) developed the following SPF based on the 'products-of-flows-to-power' relationship.

$$A = 5.59 * 10^{-6} F_1^{0.42} F_2^{0.51} k^{-2.95} \quad \dots[3.18]$$

A - Expected daily accident frequency (within 30m of intersection).

*F*₁ - major road traffic flow (veh/day).

*F*₂ - Minor road traffic flows (veh/day).

k - Dispersion parameter from Negative Binomial regression.

Hauer et al. (1988) identified 15 accident patterns in which two vehicles at an intersection can be involved and developed a SPF for each of the accident patterns. Accidents in each pattern are defined by the manoeuvres of the vehicles before the collision. These accident patterns are shown in Figure 3.9.

The form of the model equations was chosen after an exploratory analysis of the data to determine how accidents would depend on the contributory traffic flows. The guiding principle was the wish to ensure a satisfactory fit with parsimony of parameters and without violation of the obvious logical requirements.

The results of the study by Hauer et al. (1988) are given in Table 3.3.

The study by Hauer et al. (1988) has shown the benefit of using disaggregated accident data to relate these to the traffic flows to which the colliding vehicles belong. They conclude that the logic of attempting to seek an aggregate relationship between accident frequencies and some function of all flows to be unsatisfactory. They also question the suitability of the customary categorisation of accidents by initial impact (rear-end, angle, turning movement, sideswipe etc.) One cannot assume, for example, that classification of an accident as a right-angle collision implies that the vehicles were travelling at right angles to each other or that most accidents involving left or right turning will be classified as turning accidents.

Table 3.3 : Regression parameters of SPF's

Pattern	Model Form	A	b	c	K
1	$A = a(F)$	$0.2052 * 10^{-6}$			4.59
2	$A = a(F)$	$0.1014 * 10^{-6}$			1.97
3	$A = a(F_2)^c$	$8.6129 * 10^{-9}$		1.0682	1.20
4	$A = a(F_2)^c$	$8.1296 * 10^{-6}$		0.3662	5.51
5	$A = a(F_1)^b(F_2)^c$	$0.3449 * 10^{-6}$	0.1363	0.6013	1.2
6	$A = a(F_1)(F_2)^c$	$0.0418 * 10^{-6}$		0.4634	2.1
7	$A = a(F_1)^b(F_2)^c$	$0.2113 * 10^{-6}$	0.3468	0.4051	1.2
8	$A = a(F_2)^c$	$2.6792 * 10^{-6}$		0.2476	1.2
9	$A = a(F_1)^b$	$6.9815 * 10^{-9}$	1.4892		1.2
10	$A = a(F_2)^c$	$5.5900 * 10^{-12}$		2.7862	1.2
11	$A = a(F_1)^b(F_2)^c$	$1.3012 * 10^{-9}$	1.1432	0.4353	1.2
12	$A = a(F_1)^b(F_2)^c$	$0.0106 * 10^{-6}$	0.6135	0.7858	1.2
13	$A = a(F_1)^b(F_2)^c$	$0.4846 * 10^{-6}$	0.2769	0.4479	1.2
14	$A = a(F_1)^b(F_2)^c$	$1.7741 * 10^{-9}$	1.1121	0.5467	1.2
15	$A = a(F_1)^b$	$0.5255 * 10^{-6}$	0.4610		1.2

k – Dispersion parameter of Negative Binomial regression.

Models were also developed for AM, PM and off-peak conditions for patterns 1, 2, 4 and 6.

Table 3.4 show a number of macroscopic models developed by Mountain and Fawaz (1996) for priority junctions, traffic signals and roundabouts in the UK. The k value for all three models is 1.65.

Table 3.4 : Safety Performance functions

Junction type	Model
Major-minor priority	$\mu = 0.141 t_1^{0.64} t_2^{0.24}$
Traffic signals	$\mu = 0.180 t_1^{0.64} t_2^{0.24}$
Roundabouts	$\mu = 0.168 t_1^{0.64} t_2^{0.24}$

t_1 – major road flow (veh/day) and t_2 – minor road flow (veh/day)

μ - Accident frequency

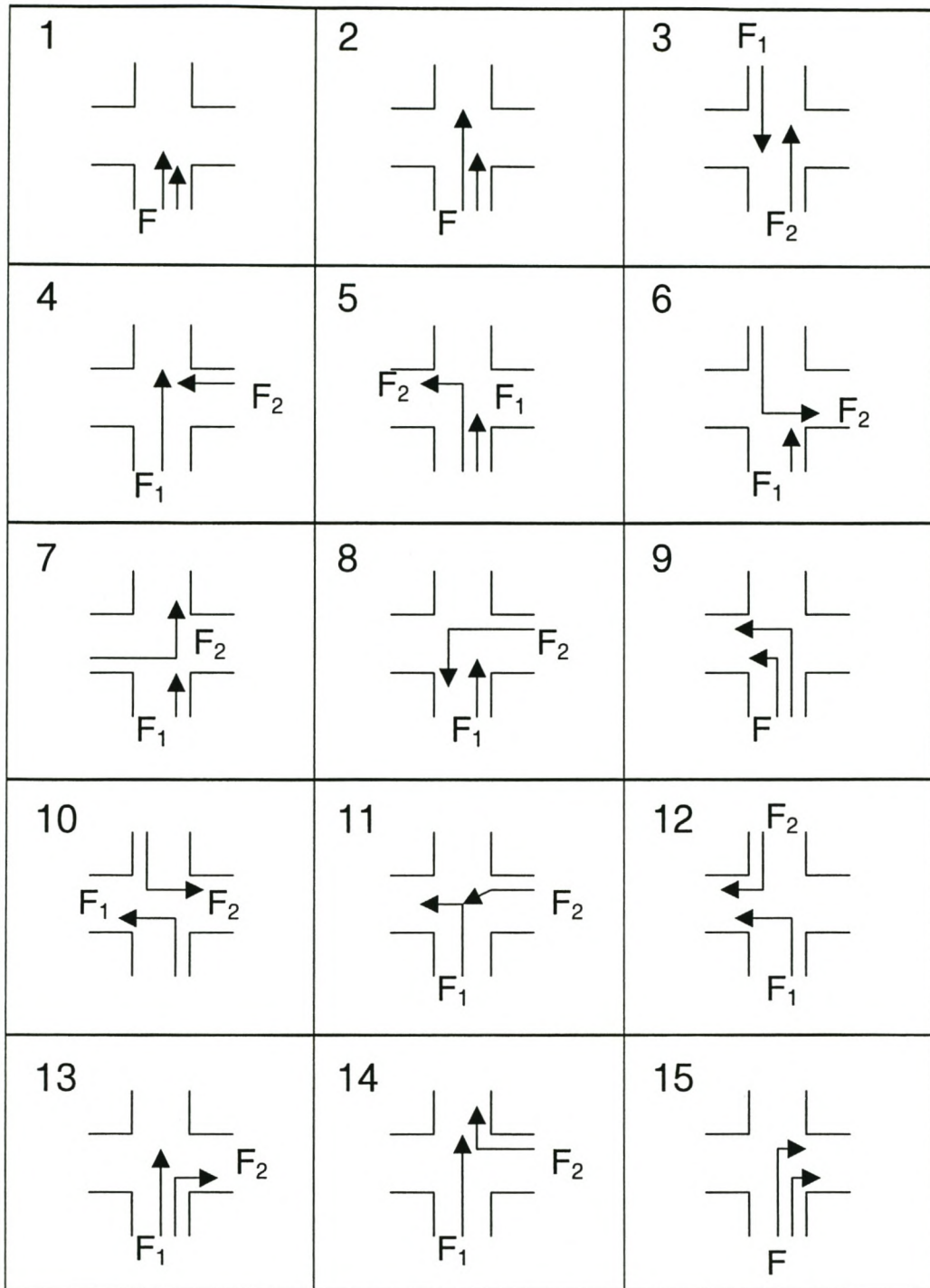


Figure 3.9 : Intersection accident patterns (Hauer et al. ; 1988) [For traffic driving on the right hand side of the road.]

3.7.3 THE ARGUMENT AVERAGING PROBLEM

Some of the Safety Performance Functions presented in the preceding section were estimated from using estimates of average traffic flow. According to Mensah and Hauer (1998) the ideal is for SPF's to represent the cause-effect relationship between accidents and the actual flows at the time of the accidents. In practice, mainly because of a lack of sufficient data, average flows are used for estimation of the SPF and not the actual flows at the time of the accidents. Relating accidents with average traffic flows forms an indirect causal link in the sense that the effect is observed over a long period of time during which the cause has assumed widely different values, but for which only the average value is known. This is referred to as the 'argument averaging problem'

To examine the issue of argument averaging consider Figure 3.10 which shows a non-linear SPF where the accident frequency μ is a function of the average flow q i.e. $\mu = \mu(q)$

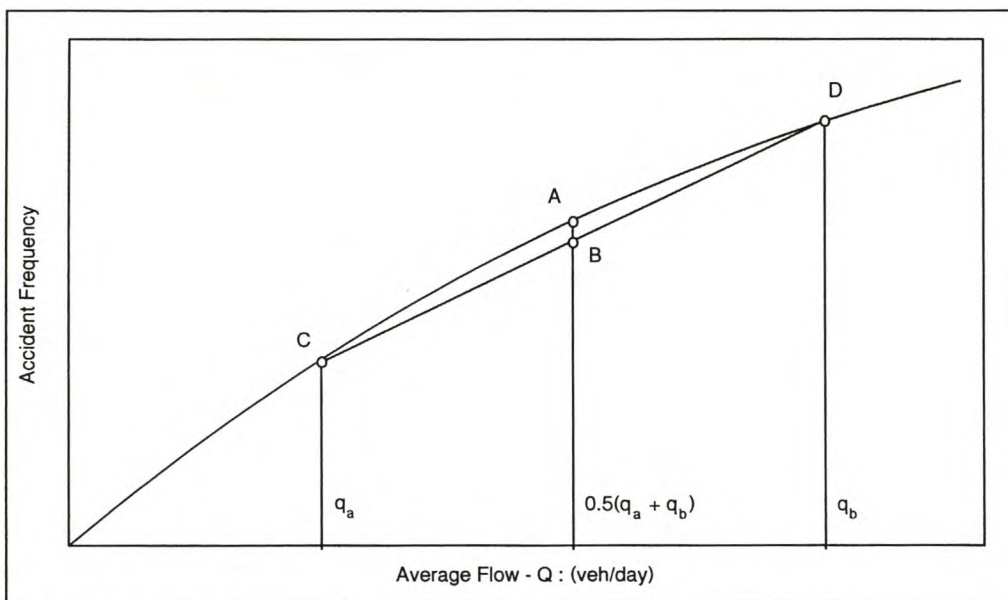


Figure 3.10 : The argument averaging problem.

If during one half of a unit of time the flow $q = q_a$ and the other half the flow $q = q_b$, then the average flow is given by $0.5(q_a+q_b)$.

Were the flow to be equal to $0.5(q_a+q_b)$ for the whole unit of time one would expect $\mu[0.5(q_a+q_b)]$ accidents per unit time. (Point A in Figure 3.10).

Since the flow q_a was for half of the time, and flow q_b for half of the time, one should expect $[\mu(a) + \mu(b)]/2$ accidents per unit time. (Point B in Figure 3.10).

If the AADT is used a function is fitted through points such as B, instead of finding the 'true' function that would pass through point A.

According to Mensah and Hauer (1998) one should not be concerned about errors due to argument averaging when :-

- a) the flow during the averaging period is constant, and
- b) when the SPF is linear for flows during the averaging period.

These conditions will arise when the period of averaging is short enough that traffic flows in it can be thought of as nearly constant and if the flow is not constant then when the segment of the SPF between the largest and the smallest flows in the period is sufficiently close to a straight line.

According to Mensah and Hauer (1998) considering what is generally known about how traffic varies over time and the likely shape of the SPF, one should generally not be concerned about averaging of periods of about an hour and that periods of longer than an hour need to be considered carefully.

Mensah and Hauer (1998) has developed a methodology to determine a correction factor w where, from Figure 3.10 :

$$w = \frac{A}{B} \qquad \dots[3.19]$$

The following information is required to determine the value of w :

- a) The distribution of traffic flows, in say hourly intervals over the study period. Information of this kind can be obtained from permanent counting stations. This information is required to estimate the CV (coefficient of variation) where :

$$CV = \frac{\sigma_q}{E(q)} \quad \dots[3.20]$$

Where :

- $\sigma(q)$ - *The standard deviation of all hourly flows across the study period.*
 $E(q)$ - *The average hourly traffic flow over the study period.*
 $\mu[E(q)]$ - *The value obtained after substituting $E(q)$ into the SPF.*

- b) The type and form of the SPF and its parameters.

Mensah and Hauer (1998) presented the following general formula to determine w from any SPF based on a single flow measure.

$$w = 1 + \frac{1}{2} \left[\frac{d^2 \mu(q)}{dq^2} \Big|_{E(q)} \right] \frac{\text{var}(q)}{\mu(E\{q\})} \quad \dots[3.21]$$

The section in the square brackets is determined by substituting $E(q)$ into the second derivative of the SPF. A detailed derivation of this equation is contained in the paper by Mensah and Hauer (1998).

The above equation leads to several observations (Mensah and Hauer ; 1998).

- i) When over the averaging period the flow varies little, that is $\text{var}(q) \rightarrow 0$ then $w = 1$ and no correction is required.

- ii) When $\mu(q)$ is constant, i.e. when the SPF is a straight line then the second derivative is zero and $w = 1$ and no correction is required.
- iii) For these values of q for which the slope of $\mu(q)$ is decreasing $w < 1$ i.e. the true level of safety will be underestimated. For values of q where $\mu(q)$ is increasing the true level of safety will be overestimated.

The following equations are presented by Mensah and Hauer (1998) to determine the value of w for each of the 4 different types of SPF's mentioned in Section 3.7.1.

a) Type 1 : The Exponential SPF

$$w = 1 + 0.5(\beta^2 - \beta)(cv)^2 \quad \dots[3.22]$$

It is evident that the correction factor in this case depends only on the exponent β and the square of the *coefficient of variation* (cv) of traffic flow.

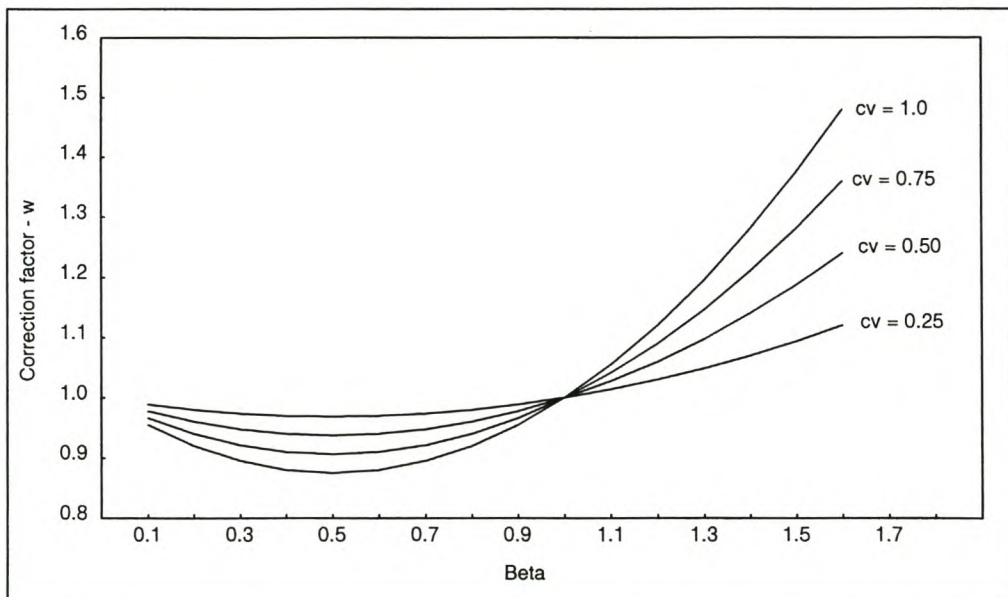


Figure 3.11 : Correction factor (w) – Exponential SPF

b) Type 2 : The Quadratic SPF

$$w = 1 + \frac{1}{\frac{\beta\alpha}{E(q)} + 1} \cdot (cv)^2 \quad \dots[3.23]$$

From Figure 3.12 its is evident that the bias can be very large at high traffic flows q and high values of cv .

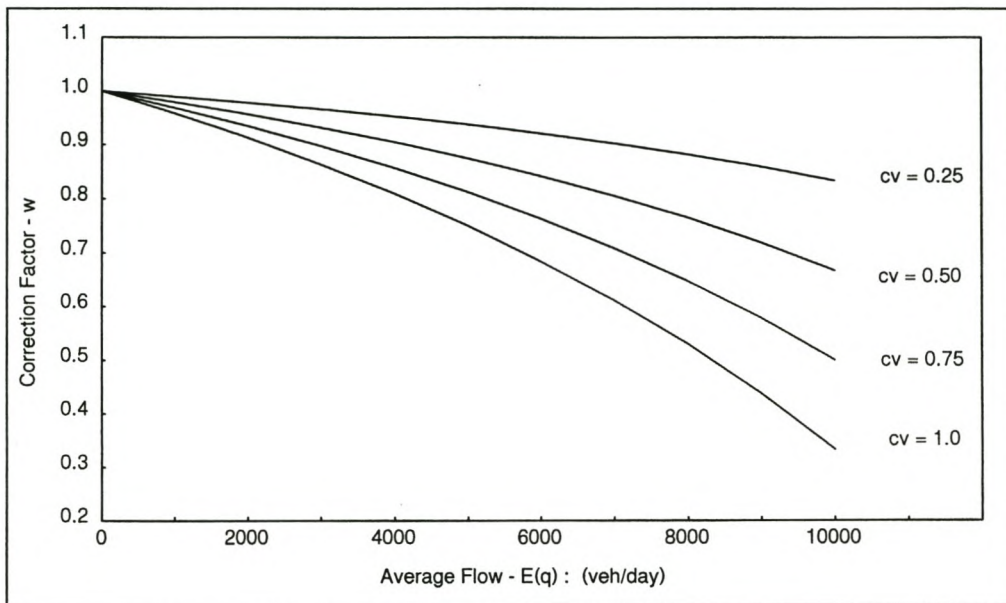


Figure 3.12 : Correction factor (w) – Quadratic SPF.

c) Types 3 and 4 : Hoerl’s Function ($k = 1$ and $k = 2$)

The correction factor w for Hoerl’s function can be estimated from Eqn. 3.24.

$$w = 1 + 0.5[\beta^2 E(q)^2 + 2k\beta E(q) + k(k - 1)](cv)^2 \quad \dots[3.24]$$

Figures 3.13 and 3.14 show how the value of w varies with different flows and *coefficient-of-variation* values.

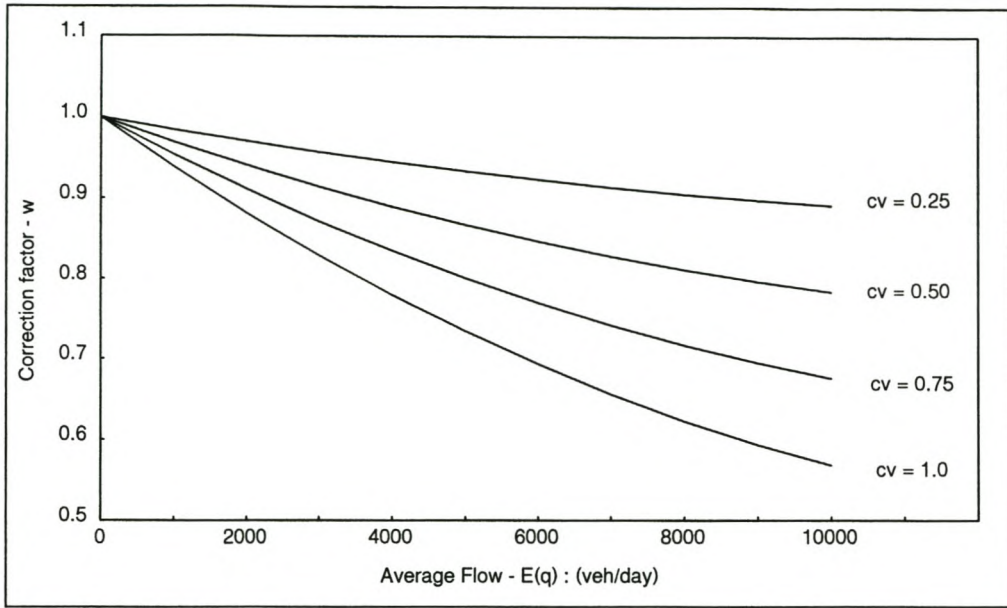


Figure 3.13 : Correction factor (w) : Hoerl's function : k = 1

It is evident from Figures 3.13 and 3.14 that the bias due to *the argument averaging problem* when using Hoerl's function could be very large at large traffic flows with a high *coefficient-of-variation*.

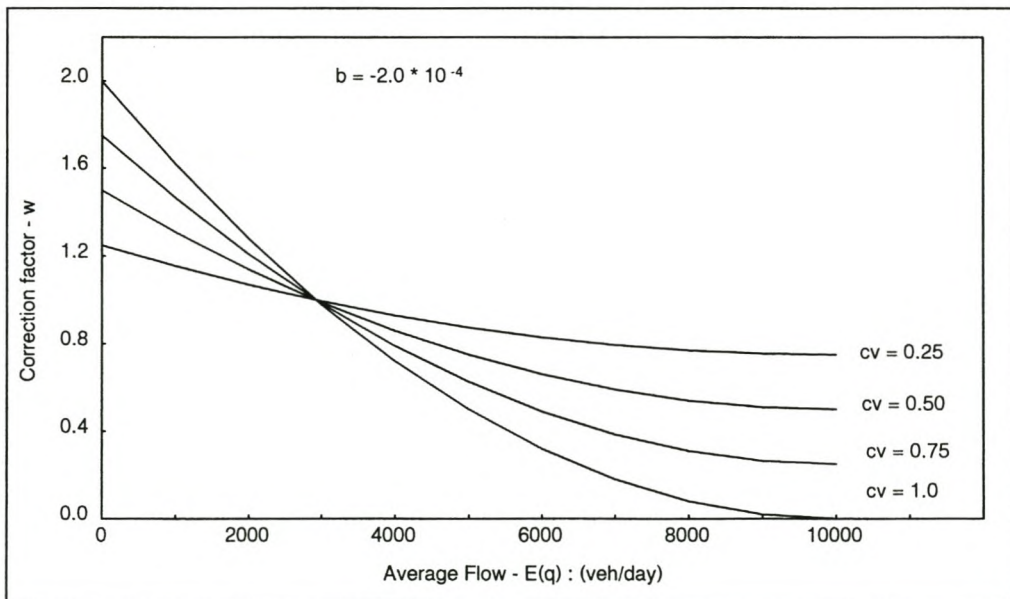


Figure 3.14 : Correction factor (w) : Hoerl's function : k = 2.

The bias introduced by developing SPF's using average traffic flows can be very large. The magnitude of this bias depends on (Mensah and Hauer ; 1998) :-

- a) The coefficient of variation (cv).
- b) The form of the SPF.
- c) The parameters of the SPF.
- d) The magnitude of the average flow (AADT).

This bias can be removed by dividing the accident counts by w before fitting the model using regression techniques. Since the determination of an appropriate value for w depends on the form of the SPF and its parameters the process must be iterative in nature. (Mensah and Hauer ; 1998)

3.7.4 FUNCTION AVERAGING PROBLEM

It is often assumed that one SPF prevailed during the entire study period T . It is however possible that during T several SPF's may apply. For example, it has been previously shown by Persaud and Musci (1995) that there exist at least 4 different SPF's for 2-lane rural roads in Ontario, Canada : 1) Single vehicle day-time, 2) Single vehicle : night-time, 3) Multi-vehicle : day-time and 4) Multi-vehicle : night-time. Different SPF's could exist for different seasons of the year, for afternoon and morning peaks etc. The problem of fitting a single SPF to data where there exist 2 or more distinct SPF's is referred to as the *function averaging problem*. (Mensah and Hauer ; 1998) This problem is illustrated in Figure 3.15.

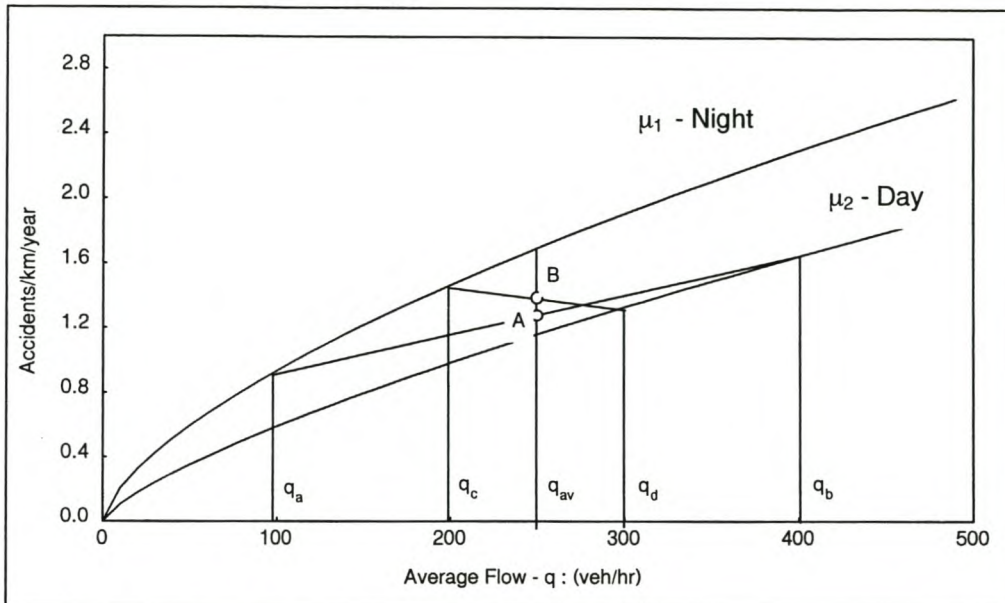


Figure 3.15 : The Function Averaging Problem

It is assumed that μ_1 represent the SPF for night-time conditions and μ_2 the SPF for daytime conditions.

With a night-time flow of q_a one would expect $\mu(q_a)$ accidents, and with a day-time flow of q_b one would expect $\mu(q_b)$ accidents. If a flow of q_a prevailed for 12 hours and a flow of q_b prevailed for 12 hours the average traffic flow is $0.5(q_a + q_b)$ and the expected accident frequency is $0.5(\mu_a + \mu_b)$. (Given by point A in Figure 3.15).

Assuming now the same amount of daily traffic ($q_a + q_b = q_c + q_d$) but a different ratio of night-time to day-time traffic it can be shown that the expected number of accidents = $0.5(\mu_c + \mu_d)$. (given by by Point B in Figure 3.15).

It is evident that even though in both cases the AADT is the same the expected number of accidents are different because of the different night- to day traffic ratios.

EXAMPLE 3.1

Using the models developed by Persaud and Musci (1995) and assuming a average day-time flow of $q_d = 2000$ and an average night-time flow of $q_n = 500$, the estimated accident numbers are shown in Table 3.5.

Table 3.5 : Expected accident numbers

Type	Day	Night	Sum	24-Hours	% Difference
Single	0.83	0.88	1.71	1.56	8.8
Multi	2.98	0.51	3.48	3.31	4.9
Sum	3.81	1.39	5.19	4.87	6.2
All	2.71	1.33	4.04	3.63	10.1
% Difference	28.9	4.3	22.2	25.5	

The total expected number of day-time single vehicle accidents is 0.83 and the total expected number of night-time accidents is 0.88. The total expected number of single vehicle accidents during the day and night is given by the sum of 0.83 and 0.88 = 1.71. Using an aggregated SPF for single vehicle accidents over a 24-hour period the expected number of accidents is 1.56. Using an aggregated SPF in this case would have underestimated the true number of accidents by approximately 8.8 %.

The expected accident frequency using an SPF based on all accidents and 24-hour flows = 3.63. The 'true' expected number of accidents is given by the sum of the expected frequencies of the 4 different SPF's which is = 5.19. Using an aggregated SPF based on all accidents and total daily flows (as is the customary practice) the 'true' accident frequency will be underestimated by 30 %.

3.8 SUMMARY AND CONCLUSION

It has been shown that the conventional measure to express the safety of a transportation system i.e. the *accident rate*, is not a suitable measure, primarily due to the fact that a non-linear relationship exists between accidents and the most common measures of exposure for road segments and intersections.

An alternative approach that relies on Safety Performance Functions was proposed. Safety Performance Functions allow the reliable comparison of

safety at equal levels of exposure. Examples of Safety Performance Functions, suitable to estimate safety on road segments, signalised and un-signalised intersections were presented.

Due to the fact that often average traffic flows are used in the estimation of Safety Performance Functions they are susceptible to the *Argument Averaging Problem* which could introduce a large and significant bias in the safety estimates. A methodology was presented on how to estimate the magnitude of this bias for road segment SPF's.

Using aggregated data, which consists of accident types normally associated with different levels of risk, the use of Safety Performance Function is also susceptible to the Function Averaging Problem which could cause the level of safety either to be over or underestimated. This problem necessitates the use of different SPF's to define the safety of an entity such as an intersection which may experience many different accident patterns, each of which are associated with a different level of risk.

CHAPTER 4

SAFETY MEASUREMENT METHODOLOGIES

4.1 INTRODUCTION

The objective of this Chapter is to provide methodologies to obtain reliable safety estimates at site and group-of-sites level taking into consideration some of the issues identified in the previous chapter.

According to Nembhard and Young (1995) there are two commonly used methods for providing safety estimates of entities a) the site-specific historical rate method, and b) the generic class method.

These two methods will be investigated and the disadvantages of each of will be identified. An alternative methodology, based on the Empirical Bayesian approach, will be presented, to show how these two methods can be combined to overcome the shortcomings of each.

The Empirical Bayesian approach and the various methodologies within this approach will be discussed in detail. Comparisons will be drawn between this approach and the conventional historical rate method.

Throughout the Chapter use will be made of the experimental data contained in Appendices A1 and A2 to illustrate some of the concepts and the differences in the accuracy and performance of the different methods.

4.2 HISTORICAL RATE METHOD

This method is based on the *conventional* definition of risk, where the risk of a system can be estimated by the accident rate as follows :

$$R = \frac{A}{E} \quad \dots[4.1]$$

R - Accident risk

A - Number of target accidents.

E - Total exposure over study period.

This method has the following disadvantages :

a) RANDOM ERROR

The historical method assumes that the observed accident frequency (*X*) at an entity is a reliable measure of the true level of safety at that location. In other words, it is assumed that the true level of safety (*m*) is known precisely and that this is equal to the observed rate/frequency at an entity. (Abbess et al ; 1981) This assumption is in violation of the fact that accidents are discrete sporadic events that follow the Poisson distribution around its true level of safety – *m*. (Al-Masaeid ; 1993). Hauer (1986) has shown that *X* is not a good estimate of *m*.

The difference between *X* and *m* is referred to as the random error (assuming the systematic error is equal to zero). Because of the random nature of *X* this random error could be significant. A compensating strategy to reduce the random error associated with the observation *X* is to increase the size of the observation. This can be achieved in one of two ways : a) increase the collection period i.e. the period over which data are collected and, b) include more entities in the study group if possible.

According to Nicholson (1987) a period of at least 5 years is required for X to be an reliable estimate of m . The problem with using such long periods for collecting data is the likelihood of significant changes in the true level of safety (m) over time at an entity, due to the influence of inevitable changes in the traffic, road and external environment. For example urban development patterns could cause an increase in pedestrian traffic or an increased proportion of night-time traffic. Both scenarios could lead to a decrease in safety.

In order to assess and to quantify the *historical rate method's* ability to predict the true level of safety the analysis described in Example 4.1 and Appendix A2 was applied to the experimental data in Appendix A1:

EXAMPLE 4.1

The true level of safety for each of the 1000 entities are known = m_i . These were randomly generated from a Gamma distribution with a mean of 4 accidents/year and a variance of 2.

For each entity 5 years of accident frequencies were randomly generated assuming a Poisson distribution with mean = m_i . These accident frequencies, X_{1i} , X_{2i} , X_{3i} , X_{4i} and X_{5i} were used to estimate m_i in the following manner :

Table 4.1 : Calculations

Study period (Years)	Expression
1	$\bar{m}_{1i} = X_{1i}$
2	$\bar{m}_{2i} = \bar{X}_2 = (X_{1i} + X_{2i})/2$
3	$\bar{m}_{3i} = \bar{X}_3 = (X_{1i} + X_{2i} + X_{3i})/3$
4	$\bar{m}_{4i} = \bar{X}_4 = (X_{1i} + X_{2i} + X_{3i} + X_{4i})/4$
5	$\bar{m}_{5i} = \bar{X}_5 = (X_{1i} + X_{2i} + X_{3i} + X_{4i} + X_{5i})/5$

For each entity and for each study period the degree-of-deviation (D) was computed as follows :

$$D_j = 100 * |(\bar{m}_{ji} - m_i)| / m_i \quad \text{Where } j = 1 \text{ to } 5 \text{ and } i = 1 \text{ to } 1000.$$

Continue ...

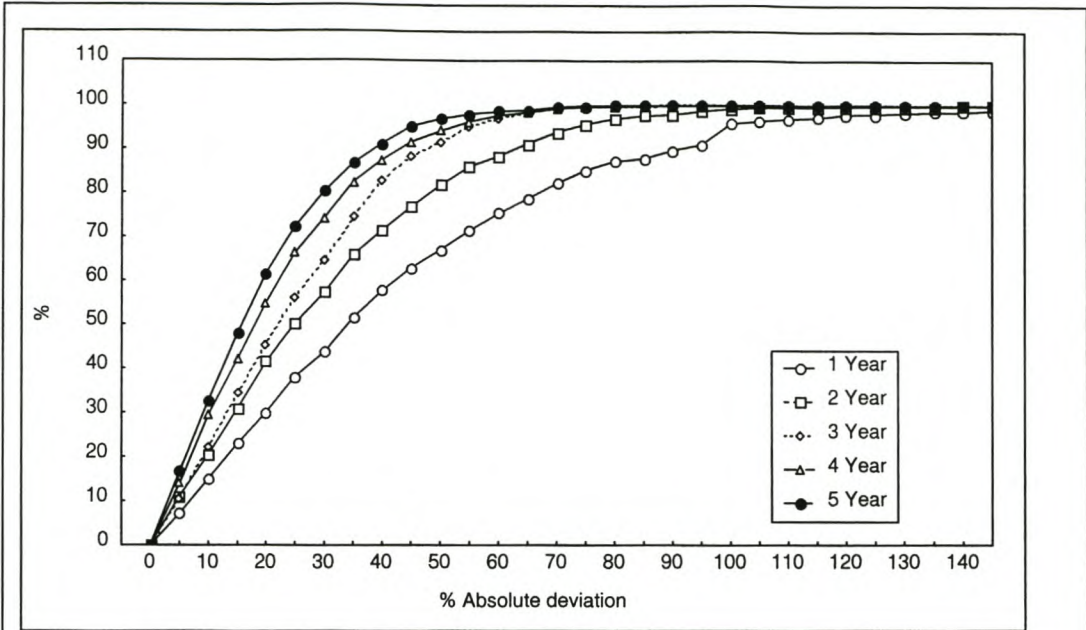


Figure 4.1 : Accuracy of different study periods

Figure 4.1 shows the cumulative density functions for the *degree-of-deviation* for the different study periods. It is evident that the accuracy of the estimates increase with increasing study periods, with the differences between successive periods gradually decreasing as the study period increase.

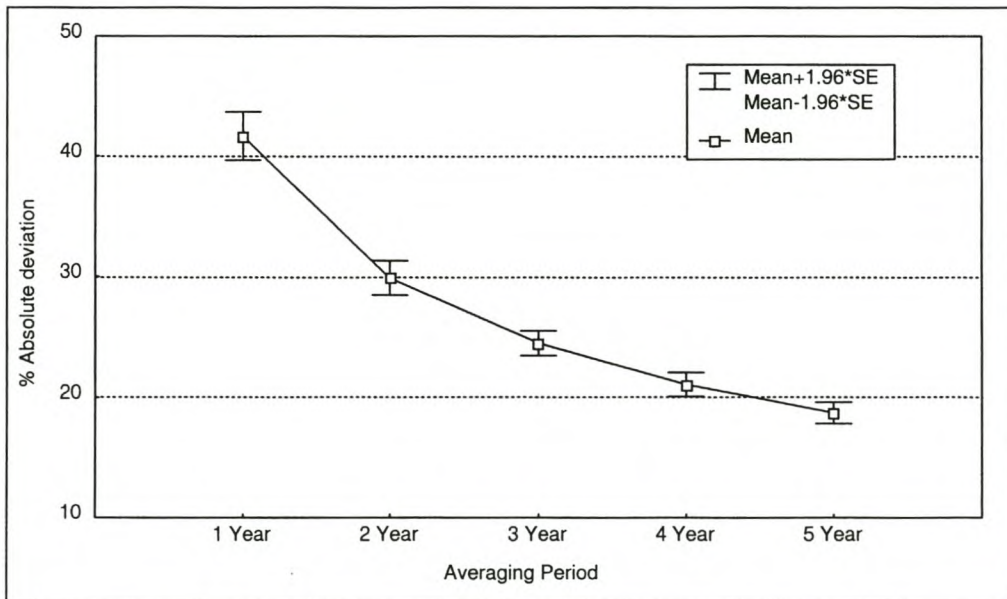


Figure 4.2 : Mean deviations for different study periods

From Figure 4.2 it is evident that even using a study period of 5 years, as recommended by Nicholson (1987), still produces an average deviation of 20 %.

b) LOW EXPOSURE

At entities with low levels of exposure zero accidents might be recorded over the study period – implying ‘perfect’ safety. This is obviously logically unsatisfactory. In addition calculating accident rates using low values of exposure could produce unstable accident rate estimates. This is because the *elasticity* of R with respect to E increases as E decreases. In other words, when E is small R could be significantly influenced by potential errors in the measurement in E , or even by small changes in E .

The elasticity of R with respect to E is given by the following equation :

$$e = -\frac{R^2}{E} \quad \dots[4.2]$$

Thus at a constant accident rate (R) the elasticity will increase with decreasing values of E .

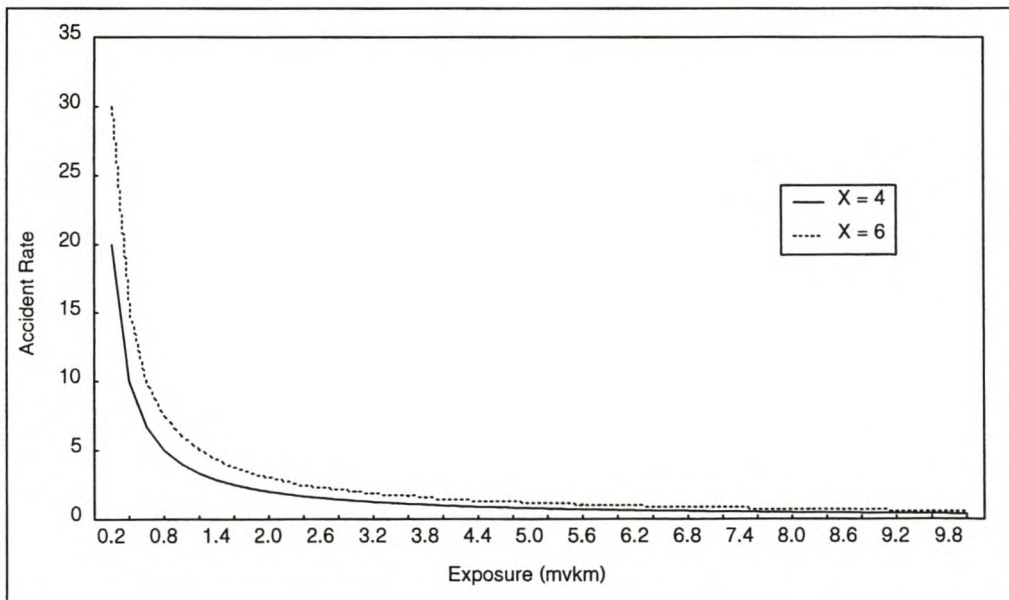


Figure 4.3 : Accident rates vs. exposure for different accident frequencies

Figure 4.3 shows how the slope of the accident rate curve increases with decreasing values of exposure.

EXAMPLE 4.2

Consider a gravel road, 10 km in length, which carries 200 vehicles per day. The average number of accidents is 1 per year. The exposure can be shown to be equal to 0.73 mvkm and the accident rate equal to 1.37 accidents per mvkm. The elasticity is equal to $-(1.37)^2/0.73 = -2.57$. In other words a 1 % error in the true value of the AADT will cause a 2.57 % error in the value of the accident rate.

Consider a road with a similar accident rate, 10 km in length and which carries 600 vehicles per day. The exposure can be shown to be equal to 2.19 mvkm and the elasticity equal to -0.86. The influence the exposure has on the accident rate when the AADT is 600 instead of 200 vehicles per day is considerably less.

Thus at very low exposure levels large accident rates can be obtained even at low accident frequencies and these accident rates can be very sensitive to changes in the level of exposure.

4.3 THE GENERIC CLASS METHOD

This method assumes that the accident risk of a location is equal to the accident risk of all locations with similar geometric, traffic and environmental characteristics – called the *reference group*. The generic accident risk is determined as follows :

$$R_i = \frac{\sum A_i}{\sum E_i} \dots[4.3]$$

The validity of the assumption, on which this method is based, is a function of the degree of similarity between a group of sites. The more similar the sites the more valid the assumption. It is however extremely unlikely that two sites are identical in every respect and that their true level of safety will be the same. In spite of being similar each site could have its own regional character,

driver population, etc. giving it a unique level of safety m . Therefore m varies from site to site. The distribution of m 's between sites in the reference group can be described by a Gamma density function with mean $E(m)$ and variance $VAR(m)$. (Al-Masaeid ; 1993)

Figure 4.4 compares the accuracy of the *generic* and the conventional *5-year historical method* in predicting the true level of safety at the 1000 entities in the experimental data contained in Appendix A. It is evident that the *historical rate method* produces more accurate estimates than the *generic* estimates.

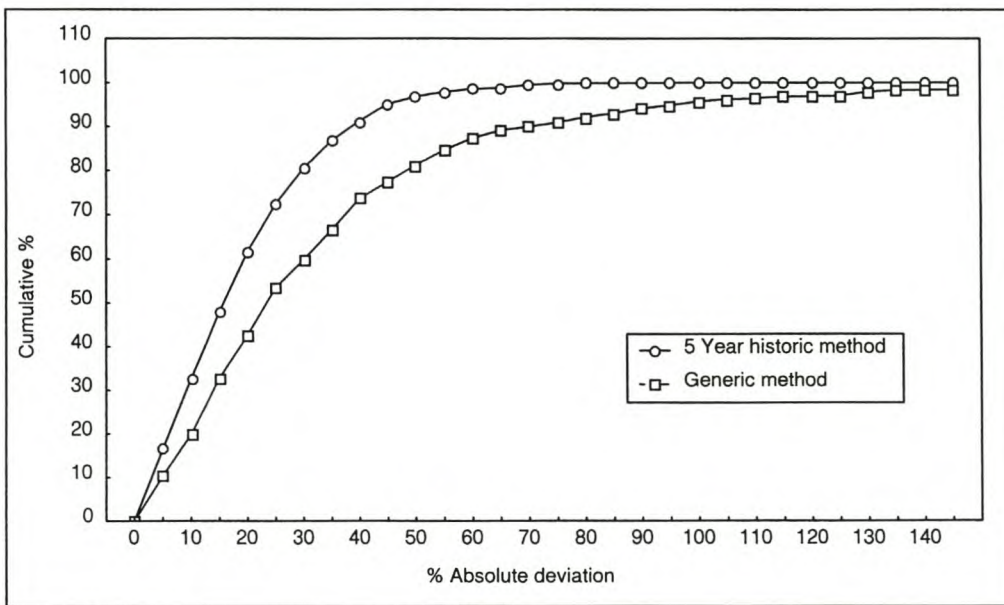


Figure 4.4 : Comparison of *generic* and *5-year historical* methods.

4.4 COMBINING THE GENERIC AND HISTORICAL RATE METHODS

According to Hauer (1997) both these methods, the historic and the generic method provide clues as to the true level of safety of an entity or a group of entities. The *generic* method provide clues as to the safety contained in all those observed and measurable characteristics that are common to all similar locations e.g. number of lanes, type of shoulders, shoulder widths etc. The *historical rate method* on the other hand provides clues as to the effect of all those unknown, unobserved and misunderstood characteristics on the level of safety at an entity i.e. those traits that make an entity absolutely unique.

According to Nemhard and Young (1995) and Hauer (1997) the true level of safety at a location is a combination of the site specific historical rate and the generic rate, where :

$$R_t = w_1 R_h + w_2 R_g \quad \dots[4.4]$$

R_t - True accident rate (level of safety).

R_h - Site specific historical accident rate.

R_g - Generic accident rate for similar sites.

w_1 and w_2 - Weighting factors where $w_1 + w_2 = 1$.

The most common methodology to estimate the weighting factors w_1 and w_2 is the empirical Bayesian methodology. (Nembhard and Young ; 1995)

4.5 THE EMPIRICAL BAYESIAN APPROACH

4.5.1 INTRODUCTION

The Empirical Bayesian approach uses *Bayes Theorem* to combine the safety estimates obtained from the *historic* and *generic* methods.

According to Al-Masaeid et al. (1993), the Bayesian approach is a probabilistic method capable of augmenting the most recent information with the available historical data or prior knowledge to achieve more accurate safety estimates.

According to Abbess et al. (1981) the Bayesian approach assumes that a probability distribution can be found before any data becomes available – this distribution is called the *prior* distribution of the parameter. Once information becomes available *Bayes theorem* can be used to convert the *prior* distribution into a *posterior* distribution. When even more information becomes available the *posterior* distribution, using *Bayes theorem*, can be updated to obtain even more accurate estimates of the parameter.

The objective is to determine the true level of safety $E(m|X)$ at a location. For this there are two clues – $E(m)$ (from the generic method) and X (the historic rate method). In order to determine the true level of safety – which is an unknown parameter, the Bayesian method requires a subjective estimate of this parameter. This is called the *prior* estimate. This estimate in itself is in all likelihood an unreliable estimate of the true level of safety at a location and has to be augmented with the observed accident experience X to obtain a more accurate and reliable *posterior* estimate, $E(m|X)$.

4.5.2 THEORETICAL FRAMEWORK

Unlike the conventional historical rate method which assumes the observed accident frequency (X) to be a constant and an appropriate measure of safety the Bayesian approach assumes the observed accident number to be variable with a Poisson probability distribution with a mean of m – where m is the ‘true’ level of safety at the location (Abbess et al. ; 1981).

$$P(x/m) = \frac{m^x e^{-m}}{x!} \quad \dots[4.5]$$

It is further assumed that m is constant over time and that the observed accident experience in different years are random variables that are Poisson distributed about m (Abbess et al. ; 1981)

The Bayesian approach further assumes that m varies between different sites and that the exact value for any particular site is unknown and is regarded as a Gamma variable with the following probability density function: (Abbess et al. ; 1981 and Higle and Witkowski ; 1988)

$$f(m) = \frac{\beta^\alpha m^{\alpha-1} e^{-\beta m}}{\Gamma(\alpha)} \quad \dots[4.6]$$

Where α , β are the parameters of the Gamma distribution.

From Ang and Tang (1975) and Abbess et al. (1981) :

$$E(m) = \frac{\beta}{\alpha} \quad \dots[4.7] \quad \text{VAR}(m) = \frac{\beta}{\alpha^2} = \frac{E(m)}{\alpha} \quad \dots[4.8]$$

The Empirical Bayesian method takes $E(m)$ and $\text{VAR}(m)$ to be the initial (*prior*) estimates of the true level of safety at an entity. In order to obtain a more reliable estimate of safety the observed accident experience, X , is combined with $E(m)$ and $\text{VAR}(m)$, using *Bayes theorem*, to obtain the more reliable *posterior* estimates, $E(m|X)$ and $\text{VAR}(m|X)$.

In order to apply *Bayes theorem* the following information is required :

- Values of $E(m)$ and $\text{VAR}(m)$.
- The Gamma parameters of $E(m)$ - α and β
- The observed accident experience - X

4.5.3 ESTIMATING $E(m)$ and $\text{VAR}(m)$

According to Hauer (1997) there are two ways to determine $E(m)$ and $\text{VAR}(m)$:

- The method of sample moments
- The multivariate regression method

Both of these methods require a suitable *reference population*.

4.5.3.1 REFERENCE POPULATION

The Bayesian approach requires a reference group to determine the *prior* distribution of m –the true level of safety at a location.

Hauer (1998) provides the following definition of what a *reference population* is :

“A reference population of entities is the group of entities that share the same set of traits as the entity in the safety of which we have an interest.”

Accident histories should not be considered in the selection process otherwise a biased estimate of m will be obtained.

Since the effect of the *prior* distribution diminishes as it is updated with observed data to form the *posterior* distribution the selection criteria for suitable reference group sites can be relaxed somewhat in order to ensure a sufficient sample size of sites.

In order to apply the multivariate regression method to estimate the safety of signalised intersections, Hauer et al. (1988) used a reference group that consisted of 145 four-legged, fixed time intersections in Metropolitan Toronto that carried two-way traffic on all approaches and which have no turn restrictions. All the intersections are on straight level sites with a speed limit of 60 km/h.

Bélanger (1994) used a reference group consisting of 149 four-legged intersections with stops on the minor approaches to estimate the safety of 4-legged unsignalised intersections. All the intersections are located in eastern Quebec. In a similar study to estimate the safety of 2-way stop controlled intersections on rural highways in Minnesota, Bonneson and McCoy (1993) used a reference group that consisted of 125 rural, four-legged, two way stop controlled intersections. In both these studies the multivariate regression method was used to determine $E(m)$ and $VAR(m)$.

To determine the safety effect of resurfacing operations on rural roads in Indiana Al-Maseied et al. (1993) used a reference group consisting of 95 undivided rural road segments to estimate $E(m)$ and $VAR(m)$ for the before and after periods using the *method of sample moments*.

4.5.3.2 THE METHOD OF SAMPLE MOMENTS

Assume there is a reference group consisting of n entities, and that each entity experienced X_{ri} accidents during the study period. The mean (μ) and standard deviation (s^2) of the observed accident frequencies are determined as follows :

$$\mu = \frac{1}{n} \sum_{i=1}^n X_{ri} \quad \dots[4.9]$$

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (X_{ri} - \mu)^2 \quad \dots[4.10]$$

According to Hauer (1997) the values of $E(m)$ and $VAR(m)$ are estimated as follows :

$$E(m) = \mu \quad \dots[4.11]$$

$$VAR(m) = s^2 - \mu \quad \dots[4.12]$$

EXAMPLE 4.3

Consider Year 1 of the experimental data contained in Appendix A.

Table 4.2 : Year 1 data – estimation of mean and variance.

X	n(X)	X*n(X)/n	(X-μ) ²
0	40	0.000	0.655
1	91	0.091	0.845
2	161	0.322	0.675
3	164	0.492	0.180
4	155	0.620	0.000
5	145	0.725	0.132
6	87	0.522	0.332
7	65	0.455	0.567
8	45	0.360	0.703
9	25	0.225	0.613
10	11	0.110	0.390
11	7	0.077	0.338
12	4	0.048	0.253
SUM	1000	4.047	5.683

$\mu = 4.047$ and $s^2 = 5.683$.

$E(m) = 4.047$ accidents/year and $VAR(m) = 5.683 - 4.047 = 1.636$

Allowing for exposure

Assume there is a reference group consisting of n entities, and that each entity experienced X_{ri} accidents during the study period and that each entity has a level of exposure equal to E_{ri} .

Let R_{ri} be the accident rate at reference entity i :

$$R_{ri} = \frac{X_{ri}}{E_{ri}} \quad \dots[4.13]$$

The mean accident rate for the reference population and its variance can be estimated as follows :

$$\bar{R} = \sum_{i=1}^n R_{ri} \quad \dots[4.14]$$

$$s_R^2 = \frac{1}{n-1} \sum_{i=1}^n (\bar{R} - R_{ri})^2 \quad \dots[4.15]$$

The harmonic mean of all the normalised traffic volumes can be determined as follows

$$\frac{1}{E'} = \frac{1}{n} \sum_{i=1}^n \frac{1}{E_{ri}} \quad \dots[4.16]$$

According to Al-Maseied et al. (1993) $E(m)$ and $VAR(m)$ can be determined as follows :

$$E(m) = \bar{R} \quad \dots[2.17]$$

$$VAR(m) = \frac{E' s_R^2 - \bar{R}}{E'} \quad \dots[2.18]$$

Calculating $E(m)$ and $VAR(m)$ using the above expressions are based on the assumption that the relationship between X and E for each entity in the reference group is linear i.e. that the Safety Performance Function is linear. If

this assumption of a linear SPF does not hold the *multivariate regression method* should be used to estimate $E(m)$ and $VAR(m)$.

4.5.3.3 THE MULTIVARIATE REGRESSION METHOD

It is often necessary to estimate the safety of entities for which a sizeable reference group population does not or cannot exist. It has been noted before that if the accident experience at a site is non-linearly related to its exposure then such a site cannot be compared to sites which do not have the same level of exposure. Therefore if the objective is, for example, to measure the safety of a section of a 2-lane rural road that carries 2000 vehicles per day the reference population should ideally also consist of 2-lane rural roads that carry 2000 vehicles per day. It is obvious that this requirement will in all likelihood result in a reference population with very few entities, if any (Hauer ; 1997).

This problem can be overcome by using multivariate regression models. Multivariate models, according to Hauer (1998), serve to 'create' an imagined reference group from which $E(m)$ and $VAR(m)$ can be determined. When a multivariate model is being fitted to accident data, it is to estimate $E(m)$ as a function of variables called 'covariates'. These 'covariates' can also be referred to as the traits or characteristics of an entity.

Hauer (1998) proposed the following procedure to estimate $E(m)$ and $VAR(m)$ using the multivariate regression method :

- 1) Develop, using data from a sufficiently large and homogenous group of reference sites, a set or sets of accident models using multivariate regression modelling techniques.

$$\mu_j = f(T_{1i}, T_{2i}, T_{3i}, T_{4i}, \dots, T_{ki}) \quad \dots[4.19]$$

- 2) For a specific site estimate the values of $E(m)$ and $VAR(m)$ as follows :

$$E_i(m) = \mu_i \quad \dots[4.20]$$

$$VAR_i(m) = \frac{[E_i(m)]^2}{k} \quad \dots[4.21]$$

k – Dispersion parameter .

The value of *k* is obtained as an ‘output’ from the regression analysis procedure. (Chapter 7)

4.5.4 APPLYING BAYES THEOREM

The objective of this section is to show how the Bayes Theorem can be used to estimate the safety of single entities and groups of entities, with and without exposure information.

4.5.4.1 SINGLE ENTITY : ACCIDENT NUMBER METHOD

According to Ang and Tang (1975) the Bayesian method can be applied to the Gamma distribution as follows :

Assume *X* accidents were observed over a period of *n* years at an entity and *E(m)* is expressed as accidents per year. The *posterior* gamma parameters are determined as follows :

$$\alpha' = \alpha + n \quad \dots[4.22]$$

$$\beta' = \beta + X \quad \dots[4.23]$$

With

$$E(m | X) = \frac{\beta'}{\alpha'} = \frac{\beta + X}{\alpha + n} \quad \dots[4.24]$$

$$VAR(m | X) = \frac{E(m | X)}{\alpha + n} \quad \dots[4.25]$$

The following equations can be derived by Substituting Equations 4.22 and 4.23 into Equations 4.24 and 4.25.

$$E(m | X) = aE(m) + (1-a)\frac{X}{n} \quad \dots[4.26]$$

$$VAR(m | X) = (1-a)E(m | X) \quad \dots[4.27]$$

where

$$a = \frac{E(m)}{E(m) + nVAR(m)} \quad \dots[4.28]$$

It is evident that Equation 4.26 is similar to Equation 4.4 with $w_1 = a$ and $w_2 = 1-a$.

According to Persaud (1993) if the variance $VAR(m|X)$ is large, i.e. when there is much unexplained variation then a will be small and the estimate will be closer to the value of X . This situation could arise when the level of exposure is quite high. For very high exposure locations therefore the observed accident count X could be a reasonably accurate safety estimate. When the amount of unexplained variation is low i.e. when $VAR(m|X)$ is small i.e. $VAR(m|X) \ll E(m|X)$ the value of a will become large and the estimate will be closer to the value of $E(m)$. This situation will typically arise when exposure levels are low.

Equation 4.26 can be rewritten as follows (Hauer ; 1997) :

$$E(m | X) = X + a[E(m) - X] \quad \dots[4.29]$$

If the observed accident frequency X is larger than $E(m)$ it means that the value of X has a large random error associated with it. In this case Equation 4.29 applies a correction factor to X i.e. $a[E(m)-X]$. Since X is larger than $E(m)$

this correction factor will be negative. Equation 4.29 serves to illustrate the process by which the Empirical Bayesian approach deals with the random error inherent in accident number observations (Hauer ; 1997).

The use of Equations 4.26, 4.27 and 4.28 to estimate the safety of a location is illustrated in Example 4.4.

EXAMPLE 4.4

Consider the experimental site with ID = 9. In the 5 year period a total of 17 accidents were recorded at this site. i.e. $X = 17$ and $n = 5$. From Example 4.3 : $E(m) = 4.047$ and $VAR(m) = 1.636$.

From Eqn 4.28 : $a = 4.047/(4.047+5(1.636)) = 0.33$

From Eqn. 4.26 : $E(m_g|X) = 0.33(4.047) + (1-0.33)(17)/5 = 3.61$ acc/year

From Eqn. 4.27 : $VAR(m_g|X) = (1-0.33)(3.61) = 2.42$

4.5.4.2 GROUP OF ENTITIES : ACCIDENT NUMBER METHOD

According to Al-Maseied et al. (1993) the expected number of accidents at a group of n similar locations is obtained by using the convolution principle as follows:

$$m_t = \sum_{i=1}^n m_i \quad \dots[4.30]$$

Where m_i has a gamma probability density function with parameters $\sum\beta'_i$ and α' .

$$\sum \beta'_i = n\beta + \sum x_i \quad \dots[4.31]$$

The expected mean and variance of m_t are :

$$E(m_t) = \frac{\sum \beta'_i}{\alpha'} \quad \dots[4.32] \qquad VAR(m_t) = \frac{\sum \beta'_i}{(\alpha')^2} \quad \dots[4.33]$$

The use of these equations to estimate the safety of a number of sites in the experimental group of data in Appendices A1 and A2 is illustrated in Example 4.5.

EXAMPLE 4.5

In Example 4.3 the prior estimates , $E(m)$ and $VAR(m)$ were determined to be equal to 4.046 and 1.636 respectively. The *prior* Gamma parameters are as follows :

$$\alpha = 4.046/1.636 = 2.47 \text{ and } \beta = (4.046)^2/1.636 = 10.0$$

X in this example represents the total accident number for a three year period i.e. $n = 3$

Table 4.3 : Groups of entities – accident number method

ID	X	β'	α'	$E(m x_i)$	$VAR(m X_i)$
1	7	17.00	5.47	3.11	1.77
2	11	21.00	5.47	3.84	2.69
3	18	28.00	5.47	5.12	4.79
4	11	21.00	5.47	3.84	2.69
5	16	26.00	5.47	4.75	4.13
6	11	21.00	5.47	3.84	2.69
7	7	17.00	5.47	3.11	1.77
8	15	25.00	5.47	4.57	3.82
9	10	20.00	5.47	3.66	2.44
10	7	17.00	5.47	3.11	1.77
TOTALS		213.000		38.94	

Using Equations 4.30 and 4.32 :

$$E(m_t) = 38.94$$

$$VAR(m_t) = 213/(5.47)^2 = 7.12$$

4.5.4.3 SINGLE ENTITY : ACCIDENT RATE METHOD

Assume the site under consideration experienced X_s accidents and has a exposure equal to E_s .

Let R_s be the accident rate at the study site :-

$$R_s = \frac{X_s}{E_s} \quad \dots[4.34]$$

Where E_s = Annual traffic in million vehicles (AADT*365/10⁶).

The estimated *prior* parameters of the gamma distribution is as follows: (Refer to Eqn. 4.17 and 4.18)

$$\alpha = \frac{E(m)}{VAR(m)} = \frac{E' \bar{R}}{E' S_R^2 - \bar{R}} \quad \dots[4.35]$$

$$\beta = \frac{[E(m)]^2}{VAR(m)} = \bar{R} \alpha \quad \dots[4.36]$$

Where E' = Harmonic mean [Eqn. 4.16], \bar{R} = Mean accident rate [Eqn. 4.14] and S_R^2 = Variance of accident rates [Eqn. 4.15].

Once the parameters of *prior* distribution have been determined, the next step is to combine the *prior* information with the site-specific data to obtain the *posterior* distribution.

$$\alpha' = \alpha + E_s \quad \dots[4.37]$$

$$\beta' = \beta + X_s \quad \dots[4.38]$$

$$E(m | X_s, E_s) = \frac{\beta'}{\alpha'} = \frac{\beta + X_s}{\alpha + E_s} \quad \dots[4.39]$$

$$VAR(m | X_s, E_s) = \frac{\beta'}{(\alpha')^2} = \frac{\beta + X_s}{(\alpha + E_s)^2} \quad \dots[4.40]$$

Substituting Equations 4.35 and 4.36 into Equations 4.39 and 4.40 will yield the following equations for $E(m|X_s, E_s)$ and $VAR(m|X_s, E_s)$:

$$E(m | X_s, E_s) = \left(\frac{E_s}{E_s + \alpha} \right) R_s + \left(\frac{\alpha}{E_s + \alpha} \right) E(m) \quad \dots[4.41]$$

$$VAR(m | X_s, E_s) = \left(\frac{E_s}{E_s + \alpha} \right) E(m | X_s, E_s) \quad \dots[4.42]$$

where

$$\alpha = \frac{E' \bar{R}}{E' S_R^2 - \bar{R}} \quad \dots[4.43]$$

where E' – the harmonic mean of the exposures (E) of all sites in the reference group.

The 'true' level of safety at a site is a combination of the observed accident rate (R) and the underlying Gamma mean, $E(m)$. As mentioned before in Section 4.2b at low levels of exposure the accident rate R becomes unstable. In Equation 4.41 the significance of R decreases with decreasing values of E while the significance of $E(m)$ increases. Thus when a site has a low level of exposure its 'true' level of safety will be closely related to the average level of safety for similar locations while the level of safety at a site with a high level of exposure will be closely related to the observed accident rate.

EXAMPLE 4.6

Appendix B1 contains accident and exposure information of a number of road sections in the Western Cape. All these road sections consist of two 3.7m lanes with 2.4 m wide surfaced shoulders.

$$E(m) = \bar{R} = 0.78 \quad S_2 = 0.84 \quad E' = 4.44$$

$$\alpha = (4.44 \cdot 0.78) / (4.44 \cdot 0.84 - 0.78) = 1.18 \text{ and } \beta = 1.18 \cdot 0.78 = 0.92$$

Consider the section of NR01004 between km 27.78 and km 54.24. During 1994 a total of 8 accidents were recorded on this section of road. During 1994 the AADT = 3085.

Continue

Example 4.6 (continue)

$$\begin{aligned}
 X_s &= 8 \\
 E_s &= (54.24-27.78)*3085 * 365*1000000 = 29.8 \text{ mvkm} \\
 R_s &= 8/29.8 = 0.269 \text{ acc/mvkm} \\
 \beta' &= 0.92 + 8 = 8.29 \text{ and } \alpha' = 1.18 + 29.8 = 30.97 \\
 E(m|X_s, E_s) &= 8.29/30.97 = 0.29 \text{ acc/mvkm} \\
 \text{VAR}(m|X_s, E_s) &= 0.29/30.97 = 0.009
 \end{aligned}$$

4.5.4.4 GROUP OF ENTITIES : ACCIDENT RATE METHOD

According to Al-Maseied et al. (1993) at the group of entities level, the total expected accident rate is given by the sum of the individual accident rates. The total expected accident rate (r_t) for a group of n sites is given by:

$$r_t = \sum_{i=1}^n r_i \quad \dots[4.44]$$

The expected value and variance of r_t is given by :

$$E(r_t) = \sum_{i=1}^n \frac{\beta'_i}{\alpha'_i} \quad \dots[4.45]$$

$$\text{VAR}(r_t) = \sum_{i=1}^n \frac{\beta'_i}{\alpha'^2_i} \quad \dots[4.46]$$

The Gamma parameters of r_t can be estimated as follows:

$$\alpha_t = \frac{E(r_t)}{\text{VAR}(r_t)} \quad \dots[4.47]$$

$$\beta_t = \frac{[E(r_t)]^2}{\text{VAR}(r_t)} \quad \dots[4.48]$$

4.5.5 PERFORMANCE

To assess the effect of different study periods on the accuracy of the Bayesian safety estimates the accident number method was used to perform the analysis in Example 4.7 and Appendix A2.

EXAMPLE 4.7

$E(m)$ and $VAR(m)$ were calculated from Year 1 data using the *method of sample moments*. Please see Example 4.3.

Using Equation 4.27 the values of m_{ij} were estimated for each of the 1000 sites for each 'collection' period from 1 to 5 years.

$$m_{ij} = aE(m) + (1-a)X_{ij}/j$$

$$a_j = E(m)/[E(m) + jVAR(m)]$$

Where $i = 1$ to 1000 and $j = 1$ to 5

X_{ij} – The sum of accidents at entity i for years 1 to j .

The next step was to compare the true level of safety (m_{ti}) with its estimate m_{ij} by using the following measure of *deviation* :

$$D_{ij} = 100*(m_{ti} - m_{ij})/(m_{ti})$$

The next step was to compile a cumulative histogram for each collection period.

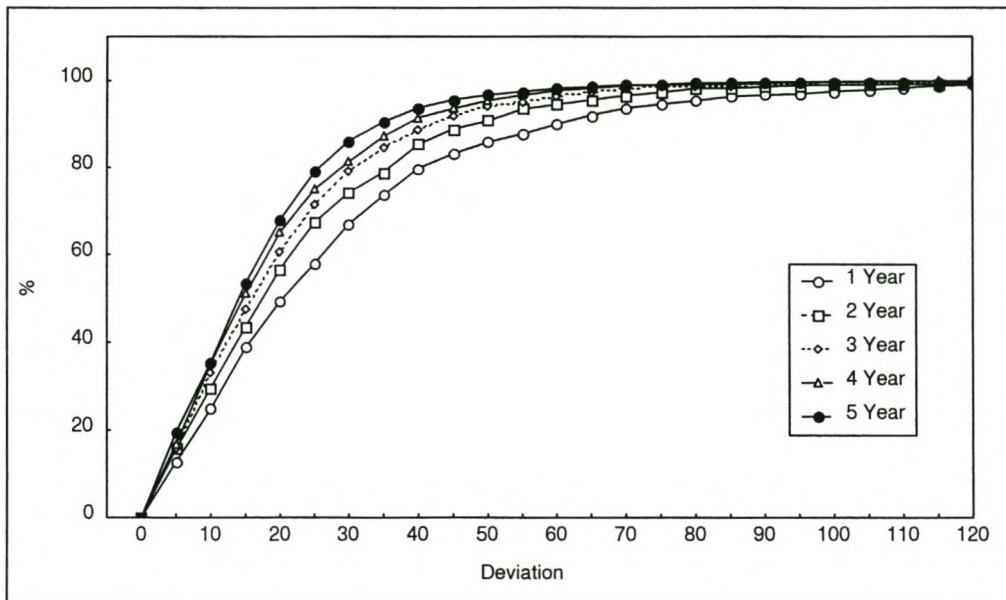


Figure 4.5 : Bayesian estimation : comparison of different study periods

Example 4.7 (continued)

Figure 4.5 shows the cumulative density functions associated with the *degree-of-deviation* for the 5 different collection periods. It is evident that as the collection period increases the estimates become more accurate, but that the rate of increase diminishes as the collection period increases.

Table 4.4 : Summary statistics : Conventional and Bayesian methods

Period	Mean	Median	Min.	Max.	Lower Quartile	Upper Quartile	Quartile Range	Std.Dev
Conventional								
1 Year	42.9	35.0	0.03	302.8	16.3	61.4	45.2	35.7
2 Year	30.0	26.3	0.03	153.6	11.7	41.8	30.1	23.2
3 Year	24.6	20.8	0.05	110.4	9.0	36.1	27.0	19.0
4 Year	21.7	18.0	0.03	97.1	8.5	31.7	23.2	16.4
5 Year	18.7	15.9	0.03	91.5	7.9	26.4	18.5	14.3
Bayesian								
1 Year	27.9	20.3	0.04	484.4	10.1	36.1	26.0	30.2
2 Year	22.9	17.3	0.02	391.4	7.7	30.4	22.7	23.9
3 Year	20.3	15.9	0.03	301.5	7.7	27.4	19.7	19.7
4 Year	18.5	14.6	0.03	239.4	7.0	25.1	18.1	17.0
5 Year	17.1	13.9	0.01	193.9	6.8	23.4	16.7	15.2

It is evident from Figure 4.6 and Table 4.4 that for similar collection periods the accuracy of the *Bayesian* estimates are consistently better than that of the *Conventional* estimates. It appears that almost the same degree of accuracy can be obtained with the 1 Year Bayesian estimates as with the 3 Year Conventional estimates.

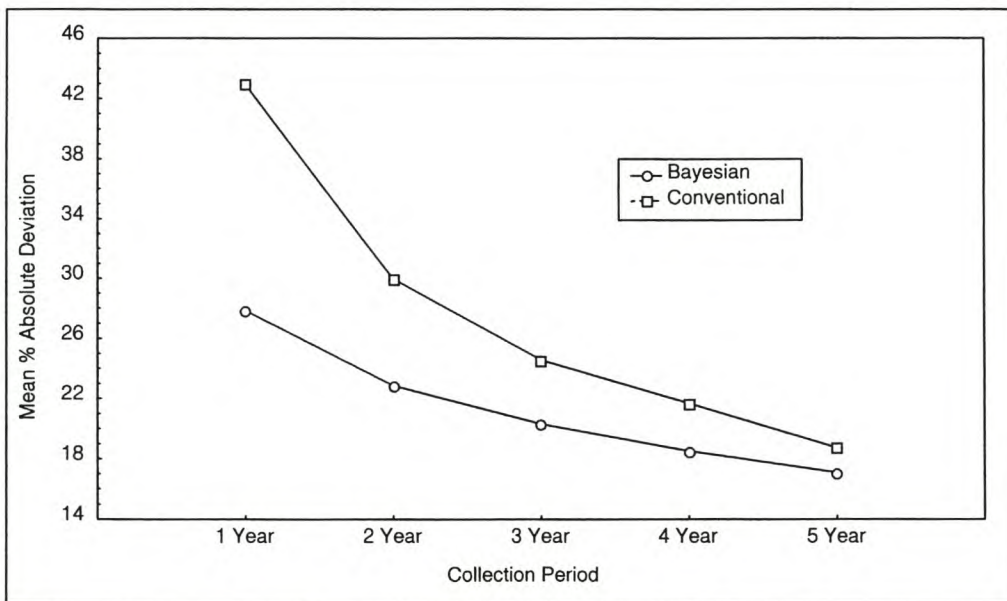


Figure 4.6 : Degree –of-deviation : Bayesian vs. Conventional methods

4.5.6 THE MULTIVARIATE REGRESSION METHOD WITHOUT A REFERENCE GROUP

According to Mountain and Fawaz (1991) there are some practical issues to consider when applying the multivariate regression method.

The determination of parameters using regression modelling requires a sufficiently large sample of reference sites in terms of measurable characteristics. If the size and homogeneity of the reference group is sacrificed it will increase bias and the variability of the parameters.

In cases where a sufficient group of reference sites is not available it is more practical to use an existing model suitable for the type of site under investigation. It is however first necessary to obtain estimates of k , the shape parameter of the assumed Gamma distribution for the between site variation in m .

If already developed models are used then k cannot be estimated directly. In such a case Mountain and Fawaz (1991) propose two possible approaches :

a) Assume an appropriate range of values for k .

This approach does not require a reference group of sites to determine $E(m)$.

According to Mountain and Fawaz (1991) a number of authors have fitted negative binomial distributions to observed accident frequencies and obtained k values in the range 0.5 to 2.8. This approach is considered not preferable because the use of even narrow ranges of k values can result in wide ranging estimates of safety.

EXAMPLE 4.8

Persaud and Musci (1995) developed the following regression model to predict the number of accidents per year per kilometre on rural 2-lane undivided roads:

$$A = 1.3392(Q)^{0.8310} \quad k = 2.9$$

A - Accidents/kilometre/year

Q - ADT (1000 vehicles)

Assume that for study road segment the AADT = 2000 i.e. Q = 2 and X = 4

$$E(m) = 1.3392(2)^{0.8310} = 2.38 \text{ acc/km/year}$$

$$\text{Var}(m) = (2.38^2)/k$$

Table 4.5 : Calculations

k	Var(m)	E(m)	A	X	E(m X)	D (%)
0.5	11.32	2.38	0.17	4	4.41	16.9
0.9	6.29	2.38	0.27	4	4.65	12.4
1.3	4.35	2.38	0.35	4	4.84	8.9
1.7	3.33	2.38	0.42	4	4.99	6.0
2.1	2.70	2.38	0.47	4	5.12	3.6
2.5	2.27	2.38	0.51	4	5.22	1.7
2.9	1.95	2.38	0.55	4	5.31	0.0

$$E(m|X) = aE(m) + (1-a)X$$

$$a = E(m)/[E(m) + \text{VAR}(m)]$$

$$D = 100*[5.31-E(m|X)]/5.31$$

The real value of k = 2.90 and hence the 'true' value of $E(m|X)$ is 5.31. D therefore measure by how much the values of $E(m|X)$ obtained at different values of k deviate from the 'true' value = 5.31

b) Estimate an appropriate value of k

Mountain and Fawaz (1991) maintain that instead of assuming a value for k a more preferable approach would be to estimate k using such data as are available about the population of sites under study. If there

are insufficient data to determine a reliable k for sites similar to the study site then a broader grouping of sites can be used to generate a sufficient number of sites.

If a broader grouping of sites is used then k can be estimated as follows: -

$$k = \phi \left(1 + \frac{1}{\theta}\right) \quad \dots[4.49]$$

Where ϕ is determined from fitting a negative binomial distribution to the observed accident frequencies of each site in the reference group. The value of ϕ can be estimated from the method of moments as follows :

$$\phi = \frac{\bar{x}^2}{s_x^2 - \bar{x}} \quad \dots[4.50]$$

The value of θ can be determined from fitting a Gamma distribution to the values of y for each site. Where y is the predicted number of accidents at a site using the prediction model.

$$\theta = \frac{\bar{y}^2}{s_y^2 - \bar{y}} \quad \dots[4.51]$$

4.5.7 THE MULTIVARIATE REGRESSION METHOD : A MORE COHERENT APPROACH

The content of this section is based exclusively on the pioneering work done by Hauer (1997) in his book *Observational Before-and-After Studies in Road Safety* (Pergamon).

According to Hauer (1997) one of the major impediments of the ‘classical’ *multivariate regression method* as discussed thus far is that a single estimate of m_i is determined for the whole analysis period. This is based on the assumption that over the analysis period the ‘true’ level of safety at a location (m_i) remains unchanged. A further impediment, is that a fixed analysis period is assumed, based on the notion that accident information prior to this period contains no useful information.

4.5.7.1 THEORETICAL FRAMEWORK

Assume there is a reference group consisting of n sites. At each site accident data and covariate values are available for each of the years $y = 1, 2, \dots, Y$.

Using multiple regression analysis a multiple regression model can be fitted to the data for each of the years $y = 1, 2, \dots, Y$.

For example :

$$E(m_{i,y}) = L_t \alpha_y Q_{i,y}^\beta \quad \dots[4.52]$$

$$VAR(m_{i,y}) = [E(m_{i,y})]^2 / k \quad \dots[4.53]$$

where

$E(m_{i,y})$ - Denote the mean of the m_i 's in year 'y' for all entities in the imagined reference population of entity i .

$VAR(m_{i,y})$ - Denote the variance of these m_i 's

- k - The dispersion parameter obtained from negative binomial regression analysis.
- $Q_{i,y}$ - Traffic flow at entity i in year y .
- L_i - The length of entity i .

In the above model the purpose of the α_y 's is to capture the influence of all factors that change from year to year, except for the change in traffic flow. The influence of changes in traffic flow is accounted for separately through Q^β . It is stated that the α_y 's can be used to account for the joint influences of year-to-year changes in weather, in economic conditions, the inclination to report accidents etc. The absence of a subscript ' i ' indicates two beliefs. First that weather, economic conditions and similar factors changed in the same manner on all road sections in the reference group. Second, that the effect of a specific change from year-to-year affects all road sections in the same manner (Hauer ; 1997).

The objective is to estimate $m_{i,y}$, the expected number of accidents of the entity i in year y , knowing that the entity experienced $X_{i,1}$, $X_{i,2}$, ... $X_{i,y}$ accidents. The different values of $m_{i,y}$ for an entity i are not independent of one another. According to Hauer (1997) certain characteristics of an entity do not change from year to year e.g. lane widths, gradient, curvature etc. Some characteristics do however change from year to year – as quantified by the α_y 's, e.g. weather, economic situation, traffic composition etc. One can therefore expect that there will be some similarity in m_i over the years, but that there will also be some change in m_i from year to year.

Hauer (1997) provides a detailed derivation of the following expressions :

$$\frac{m_{i,y}}{m_{i,1}} = \frac{E(m_{i,y})}{E(m_{i,1})} = C_{i,y} \quad \dots[4.54]$$

$$m_{i,y} = m_{i,1} C_{i,y} \quad \dots[4.55]$$

$$VAR(m_{i,1}) = C_{i,y}^2 VAR(m_{i,1}) \quad \dots[4.56]$$

The underlying assumption of these expression is that how the m_i of an entity changes from year to year can be adequately represented by the change in the covariates over time, and by the model equation. The effect of these expressions is that $m_{i,y}$ and $VAR(m_{i,y})$ are all a function of $m_{i,1}$.

Hauer (1997) provides the following expressions to estimate $m_{i,1}$ and $VAR(m_{i,1})$:

$$m_{i,1} = \frac{k + \sum_{y=1}^Y X_{i,y}}{\frac{k}{E(m_{i,1})} + \sum_{y=1}^Y C_{i,y}} \quad \dots[4.57]$$

$$VAR(m_{i,1}) = \frac{m_{i,1}}{\frac{k}{E(m_{i,1})} + \sum_{y=1}^Y C_{i,y}} \quad \dots[4.58]$$

The use of these equations to estimate the safety of an entity, as it changes from year to year, is illustrated in Example 4.9.

EXAMPLE 4.9

Consider a hypothetical segment of road with length = 2km that during a 6 year period experienced the following number of accidents - {4, 1, 3, 4, 2, 5}. Using a reference group of similar sites a total of 6 multivariate regression models were fitted to the data, one for each year, – similar in form to the model in Equation 3.5. The β parameter for all the models = 0.78. The α parameters for six different models are shown in the Table below. The k value obtained from calibrating the prediction model for year 1 = 4.5

Table 4.6 : Data and calculations

y	α_y	$Q_{i,y}$	$E(m_{i,y})$	$C_{i,y}$	$X_{i,y}$	$m_{i,y}$	$VAR(m_{i,y})$
1	0.00271	1250	1.41	1.000	4	2.49	0.264
2	0.00295	1156	1.45	1.024	1	2.55	0.277
3	0.00277	1277	1.47	1.039	3	2.59	0.285
4	0.00265	1334	1.45	1.029	4	2.56	0.280
5	0.00284	1288	1.51	1.073	2	2.67	0.304
6	0.00278	1305	1.50	1.061	5	2.64	0.297
TOTALS				6.226	19	15.5	

From Eqn. 4.57 and Eqn. 4.58:

$$m_{i,1} = (4.5 + 19)/(4.5/1.41 + 6.226) = 2.49$$

$$VAR(m_{i,1}) = 2.49/(4.5/1.41+6.226) = 0.264$$

From Eqn. 4.55 and Eqn. 4.56:

$$m_{i,2} = 2.49 (1.024) = 2.549$$

$$VAR(m_{i,2}) = (1.024)^2(0.264) = 0.277$$

4.5.8 STATISTICAL INFERENCE

Both *prior* and *posterior* Bayesian safety estimates follow a Gamma distribution.

If the parameters of the Gamma distribution are known, using distribution theory, a number of statistical inferences can be made concerning the safety estimate of an entity or a group of entities.

The following inferences could be of interest :

- The probability that the safety estimate lies between two values.
- The probability that the safety estimate lies above or below a certain value.
- The values of two-tailed or one-tailed confidence intervals.
- The probability of two Bayesian estimates being different from one another.

The probability density function of a Gamma distribution is as follows :

$$f(m) = \frac{\beta^\alpha m^{\alpha-1} e^{-\beta m}}{\Gamma(\alpha)} \quad \dots[4.59]$$

To determine the probability that m is larger than b and smaller than a the probability distribution function can be used as follows :

$$P(a < m \leq b) = \int_a^b \frac{\beta^\alpha m^{\alpha-1} e^{-\beta m}}{\Gamma(\alpha)} dm \quad \dots[4.60]$$

In Figure 4.7 the probability density function and probability distribution function of an estimate with parameters $\alpha = 1$ and $\beta = 3.4$ is shown.

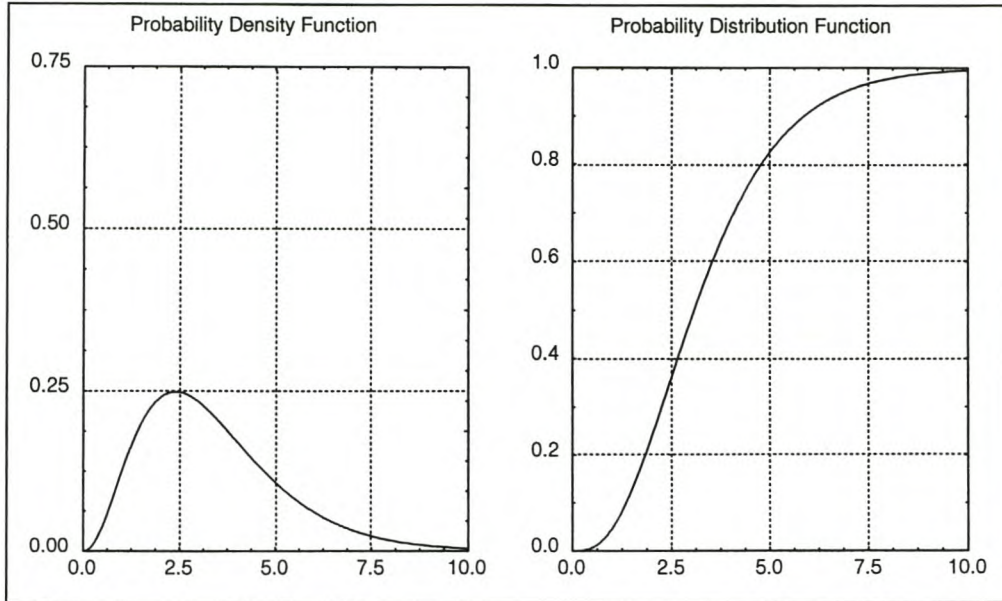


Figure 4.7: Probability density and distribution functions for a Gamma distribution with $\alpha = 1$ and $\beta = 3.4$.

The integral in Equation 4.59 can be solved using the *GAMMADIST* function of Microsoft® Excel® as shown in Table 4.7. In any Excel® function that refers to the Gamma distribution *alpha* = β and *beta* = $1/\alpha$.

Table 4.7 : Using the GAMMADIST function of Excel®

Prob.	Function
$P(m \leq b)$	<code>=GAMMADIST(b, β, 1/α, TRUE)</code>
$P(a < m \leq b)$	<code>=GAMMADIST(b, β, 1/α, TRUE) - GAMMADIST(a, β, 1/α, TRUE)</code>
$P(m > b)$	<code>=1 - GAMMADIST(b, β, 1/α, TRUE)</code>

To determine the value of m_{crit} for which $P(m \leq m_{crit}) = 0.95$ the *GAMMAINV* function of Microsoft® Excel® can be used as follows :

`=GAMMAINV(0.95, β , 1/ α)`

...[4.61]

The critical values of m at which $P(m \leq m_{crit}) = 0.95$ or $P(m \leq m_{crit}) = 0.05$ can be obtained from Tables 4.8 and 4.9 respectively. These Tables are used as follows :

- a) Using the β parameter of the Gamma distribution locate the appropriate value of G .
- b) Having obtained the G value m_{crit} is obtained using the α parameter as follows:

$$m_{crit} = G/\alpha \quad \dots[4.62]$$

EXAMPLE 4.10

In Example 4.6 the following estimates were obtained for a segment of road :

$$E(m|X_s, E_s) = 8.29/30.97 = 0.29 \text{ acc/mvkm}$$

$$\text{VAR}(m|X_s, E_s) = 0.29/30.97 = 0.009$$

The Gamma parameters are :

$$\beta' = 0.92 + 8 = 8.92$$

$$\alpha' = 1.18 + 29.8 = 30.98$$

To determine the m_{crit} such that $P(m \leq m_{crit}) = 0.05$ obtain from Table 4.9 the value of G where $\beta = 8.9$.

$$G = 4.84 \text{ and } m_{crit} = 4.84/30.97 = 0.156.$$

To determine m_{crit} such that $P(m \leq m_{crit}) = 0.95$ obtain from Table 4.8 the value of G where $\beta = 8.9$.

$$G = 14.32 \text{ and } m_{crit} = 14.32/30.97 = 0.462.$$

The 90 % confidence interval of $E(m|X,E)$ is therefore $\{0.156 ; 0.462\}$.

Al-Masaeid et al. (1993) provides the derivation for the following expressions that can be used to determine the probability that two Bayesian estimates are different from each other :

Assume there are two locations with Bayesian estimates m_1 and m_2 respectively. The Gamma parameters of Site 1 are α and β while the Gamma parameters of Site 2 are α' and β' .

Assuming that m_1 and m_2 are independent the joint probability density function is given by :

$$f(m_1, m_2) = \frac{\beta^\alpha m_1^{\alpha-1} e^{-\beta m_1}}{\Gamma(\alpha)} \cdot \frac{\beta'^{\alpha'} m_2^{\alpha'-1} e^{-\beta' m_2}}{\Gamma(\alpha')} \quad \dots[4.63]$$

$P(m_1 < m_2)$ can be estimated as follows :

$$p(m_1 < m_2) = 1 - \sum_{j=0}^{\alpha-1} \left[\frac{\beta}{\beta'} \right] \left[\frac{\beta'}{\beta + \beta'} \right]^{\alpha+j} \cdot \frac{\Gamma(\alpha'+j)}{\Gamma(j+1)\Gamma(\alpha')} \quad \dots[4.64]$$

The Eqn. 4.64 can be solved using a spreadsheet. E.g. the Gamma function $\Gamma(\alpha')$ can be solved using a combination of the EXP and GAMMALN functions in Microsoft® Excel®. $\Gamma(\alpha') = EXP(GAMMALN(\alpha'))$.

EXAMPLE 4.11

Two sites have the following parameters:

Site 1 $\alpha = 4.0, \beta = 1.9, E(m_a|x_a) = 2.10$

Site 2 : $\alpha' = 3.4, \beta' = 1.2, E(m_b|x_b) = 2.80$

$$\text{If } D_j = \left[\frac{\beta}{\beta'} \right] \left[\frac{\beta'}{\beta + \beta'} \right]^{\alpha+j} \cdot \frac{\Gamma(\alpha'+j)}{\Gamma(j+1)\Gamma(\alpha')}$$

Then $D_0 = 0.46$ and $D_1 = 0.25$

$$\sum_{j=1}^{0.9} D_j = D_0 + 0.9D_1 = 0.46 + 0.9(0.25) = 0.685$$

$$P(m_1 < m_2) = 1 - 0.685 = 0.315$$

Table 4.8 : Incomplete Gamma function values ($\alpha = 1$) : 95 % Degree of Confidence

β	G	β	G	β	G	β	G
0.1	0.58	4.1	7.90	8.1	13.28	12.1	18.33
0.2	1.03	4.2	8.04	8.2	13.41	12.2	18.46
0.3	1.37	4.3	8.18	8.3	13.54	12.3	18.58
0.4	1.66	4.4	8.32	8.4	13.66	12.4	18.70
0.5	1.92	4.5	8.46	8.5	13.79	12.5	18.83
0.6	2.16	4.6	8.60	8.6	13.92	12.6	18.95
0.7	2.38	4.7	8.74	8.7	14.05	12.7	19.07
0.8	2.60	4.8	8.88	8.8	14.18	12.8	19.20
0.9	2.80	4.9	9.02	8.9	14.31	12.9	19.32
1.0	3.00	5.0	9.15	9.0	14.43	13.0	19.44
1.1	3.19	5.1	9.29	9.1	14.56	13.1	19.57
1.2	3.37	5.2	9.43	9.2	14.69	13.2	19.69
1.3	3.55	5.3	9.57	9.3	14.82	13.3	19.81
1.4	3.73	5.4	9.70	9.4	14.94	13.4	19.93
1.5	3.91	5.5	9.84	9.5	15.07	13.5	20.06
1.6	4.08	5.6	9.97	9.6	15.20	13.6	20.18
1.7	4.25	5.7	10.11	9.7	15.33	13.7	20.30
1.8	4.42	5.8	10.24	9.8	15.45	13.8	20.42
1.9	4.58	5.9	10.38	9.9	15.58	13.9	20.55
2.0	4.74	6.0	10.51	10.0	15.71	14.0	20.67
2.1	4.91	6.1	10.65	10.1	15.83	14.1	20.79
2.2	5.06	6.2	10.78	10.2	15.96	14.2	20.91
2.3	5.22	6.3	10.91	10.3	16.08	14.3	21.03
2.4	5.38	6.4	11.05	10.4	16.21	14.4	21.16
2.5	5.54	6.5	11.18	10.5	16.34	14.5	21.28
2.6	5.69	6.6	11.31	10.6	16.46	14.6	21.40
2.7	5.84	6.7	11.45	10.7	16.59	14.7	21.52
2.8	5.99	6.8	11.58	10.8	16.71	14.8	21.64
2.9	6.15	6.9	11.71	10.9	16.84	14.9	21.77
3.0	6.30	7.0	11.84	11.0	16.96	15.0	21.89
3.1	6.44	7.1	11.97	11.1	17.09	15.1	22.01
3.2	6.59	7.2	12.11	11.2	17.21	15.2	22.13
3.3	6.74	7.3	12.24	11.3	17.34	15.3	22.25
3.4	6.89	7.4	12.37	11.4	17.46	15.4	22.37
3.5	7.03	7.5	12.50	11.5	17.59	15.5	22.49
3.6	7.18	7.6	12.63	11.6	17.71	15.6	22.61
3.7	7.32	7.7	12.76	11.7	17.84	15.7	22.73
3.8	7.47	7.8	12.89	11.8	17.96	15.8	22.86
3.9	7.61	7.9	13.02	11.9	18.08	15.9	22.98
4.0	7.75	8.0	13.15	12.0	18.21	16.0	23.10

Table 4.9 : Incomplete Gamma function values ($\alpha = 1$): 5 % Degree of Confidence

β	G	β	G	β	G	β	G
0.1	0.00	4.1	1.42	8.1	4.05	12.1	7.00
0.2	0.00	4.2	1.48	8.2	4.12	12.2	7.08
0.3	0.00	4.3	1.54	8.3	4.19	12.3	7.15
0.4	0.00	4.4	1.60	8.4	4.26	12.4	7.23
0.5	0.00	4.5	1.66	8.5	4.34	12.5	7.31
0.6	0.01	4.6	1.72	8.6	4.41	12.6	7.38
0.7	0.01	4.7	1.78	8.7	4.48	12.7	7.46
0.8	0.02	4.8	1.85	8.8	4.55	12.8	7.54
0.9	0.03	4.9	1.91	8.9	4.62	12.9	7.61
1.0	0.05	5.0	1.97	9.0	4.70	13.0	7.69
1.1	0.07	5.1	2.03	9.1	4.77	13.1	7.77
1.2	0.09	5.2	2.10	9.2	4.84	13.2	7.84
1.3	0.12	5.3	2.16	9.3	4.91	13.3	7.92
1.4	0.15	5.4	2.22	9.4	4.99	13.4	8.00
1.5	0.18	5.5	2.29	9.5	5.06	13.5	8.08
1.6	0.21	5.6	2.35	9.6	5.13	13.6	8.15
1.7	0.24	5.7	2.42	9.7	5.20	13.7	8.23
1.8	0.28	5.8	2.48	9.8	5.28	13.8	8.31
1.9	0.32	5.9	2.55	9.9	5.35	13.9	8.39
2.0	0.36	6.0	2.61	10.0	5.43	14.0	8.46
2.1	0.40	6.1	2.68	10.1	5.50	14.1	8.54
2.2	0.44	6.2	2.75	10.2	5.57	14.2	8.62
2.3	0.48	6.3	2.81	10.3	5.65	14.3	8.70
2.4	0.53	6.4	2.88	10.4	5.72	14.4	8.78
2.5	0.57	6.5	2.95	10.5	5.80	14.5	8.85
2.6	0.62	6.6	3.01	10.6	5.87	14.6	8.93
2.7	0.67	6.7	3.08	10.7	5.94	14.7	9.01
2.8	0.72	6.8	3.15	10.8	6.02	14.8	9.09
2.9	0.77	6.9	3.22	10.9	6.09	14.9	9.17
3.0	0.82	7.0	3.29	11.0	6.17	15.0	9.25
3.1	0.87	7.1	3.35	11.1	6.24	15.1	9.32
3.2	0.92	7.2	3.42	11.2	6.32	15.2	9.40
3.3	0.98	7.3	3.49	11.3	6.39	15.3	9.48
3.4	1.03	7.4	3.56	11.4	6.47	15.4	9.56
3.5	1.08	7.5	3.63	11.5	6.55	15.5	9.64
3.6	1.14	7.6	3.70	11.6	6.62	15.6	9.72
3.7	1.19	7.7	3.77	11.7	6.70	15.7	9.80
3.8	1.25	7.8	3.84	11.8	6.77	15.8	9.88
3.9	1.31	7.9	3.91	11.9	6.85	15.9	9.96
4.0	1.37	8.0	3.98	12.0	6.92	16.0	10.04

4.6 SUMMARY and CONCLUSION

In this Chapter both the *Conventional* and *Bayesian* methodologies were presented to measure the safety of road infrastructure elements. Certain shortcomings of the *Conventional* methods to road safety measurement were identified and discussed. The *Empirical Bayesian* methodology was discussed in detail and it has been shown how *Bayesian methods* can compensate for the shortcomings of the *Conventional* methods.

It has been shown from analysing the experimental data in Appendices A1 and A2 with both the *Conventional* and *Bayesian* methodologies that the latter produce results that are generally more accurate, but that the relative 'benefit' of using the *Bayesian* method decreases as the study period increases.

Certain impediments of the 'classical' Bayesian methods, as described in the Chapter, were identified and a more coherent approach, based on the work of Hauer (1997), was proposed. This approach acknowledged the fact that the level of safety at a location may change from year to year.

CHAPTER 5

THE IDENTIFICATION AND RANKING OF HAZARDOUS LOCATIONS

5.1 INTRODUCTION

A road safety remedial programme requires an effective means of identifying hazardous locations. Hazardous locations according to Deacon et al. (1997) are those locations where the accident pattern or the level of safety are abnormally severe when compared with similar locations elsewhere, and for which improvements can be made through a variety of engineering measures.

The identification of hazardous locations is the first step in a road safety remedial programme. In short the procedure is as follows :

- a) Identification of hazardous locations.
- b) Preliminary ranking of selected locations.
- c) Detailed investigation.
- d) Final ranking based on cost effectiveness.
- e) Design and implementation.
- f) Monitoring and evaluation.

Applying an identification procedure is an important first step because it would be impractical to conduct a road safety investigation on each and every location in the road network. The detailed investigation of sites as well as the design and implementation of remedial measures require the application of potentially scarce resources – money, time, equipment, personnel etc. Road authorities generally have limited budgets for the purpose of implementing road safety remedial programmes. It is therefore the road authority's responsibility and moral obligation to ensure that resources are applied in a manner, and to those sites, that will yield the maximum possible economic returns. In identifying hazardous locations it is important that the chosen

methodology selects as many of the truly hazardous locations as possible and discards as many of the non-hazardous locations as possible.

Hauer and Persaud (1984) compare the identification process to a 'sieve' that retains the hazardous locations and lets the non-hazardous locations through. The degree of *efficiency* of an identification process is defined as the degree to which it retains truly hazardous locations and discards truly non-hazardous locations.

It is possible that the identification process yields more hazardous locations than a road authority has the resources to investigate. Ranking the selected hazardous locations according to the safety benefits that could potentially be achieved at a site, would ensure that available resources are applied to those hazardous locations that would yield the best economic returns.

Methods to identify and rank hazardous locations can be divided into two categories : *Conventional* methods and *Bayesian* methods. *Conventional* methods rely on conventional safety estimates, and *Bayesian* methods rely on Bayesian safety estimates and, in some cases, Safety Performance Functions.

The objective of this Chapter is to present and to evaluate both the *Conventional* and *Bayesian* methodologies. These two methodologies will be compared in terms of their *efficiency* and ability to rank hazardous locations according to their expected economic benefits.

5.2 PROGRAMME DESIGN

Nicholson (1989) refers to 4 types of plans for accident reduction :-

a) Single site plans

- The treatment of single hazardous locations e.g. intersections, ramps, bridges, horizontal curves, short segments.

b) Route plans

- A whole route or road segments with a poor level of safety is treated with a remedial measure, such as for example, surfacing gravel shoulders, widening lanes, resurfacing etc.

c) Area action plans

- A number of remedial measures are implemented in an area that has been identified as having a particularly poor safety record. For example the implementation of various traffic calming measures (mini-circles, speedhumps, street closures etc.) in a town or suburb would constitute an area action plan.

d) Mass action plans

- A common accident problem is treated by an appropriate remedial action. An example of a mass action plan would be the wide-scale resurfacing of intersection approaches to reduce head-rear or right-angle collisions at intersections.

According to Nicholson (1989) which type of plan is most appropriate to implement can be determined by evaluating the level and extent to which accidents are clustered in the study area. A high level of accident clustering would indicate that a *Single-site* type plan would be the most appropriate, while lower levels of clustering would indicate that *Route* or *Area* type plans are more suitable.

Thus with high levels of clustering, in order to maximise the expected accident reduction and economic benefit, a road authority should commence with a *single-site* (blackspot) programme. If the programme is effective the level of clustering should decrease, after which it may be necessary to adopt a higher level of aggregation i.e. a *Route* ('blackroute') programme. Routes may contain a number of single sites which do not qualify as 'blackspots' but when considered together give 'blackroutes'.

The successful treatment of such routes will lead to a further reduction in the level of clustering to the extent that even a higher level of aggregation is required.

Nicholson (1989) provides a number of methods to quantify the level of accident clustering on a network, but fail to provide threshold values as to when the different type of plans are more appropriate. The methods proposed by Nicholson (1989) can however be used to assess whether or not there were changes in the level of accident clustering.

According to Nicholson (1989) there is a natural progression from single-site plans to area plans with the expected accident reduction and economic return declining with the 'law of diminishing returns'.

Even though single-site plans yield the best results in terms of accident reductions and economic benefits relative to the other type of plans it is not always the most appropriate course of action to take. Nicholson (1989) reports that since the mid 80's systematic accident reduction efforts in New Zealand concentrated primarily on single-site plans, but that subsequently analysis has shown this to be the wrong approach because of the low levels of clustering that already existed at the time.

The identification and ranking in this chapter are aimed specifically at *single-site* and *route* action plans.

5.3 COMPONENTS OF IDENTIFICATION METHODS

An important consideration when attempting to identify hazardous locations is how to define *single-sites* and *routes*.

5.3.1 SINGLE SITES

According to Deacon et al. (1975) several considerations are paramount to determining the appropriate spot length :

- a) The spot length can be no smaller than the minimum distance increment for reporting accident locations. If accidents are reported to the nearest 100m then the spot length cannot be less than 100m.
- b) The spot length should influence errors that will occur in reporting accident locations. Such errors are inevitable partly because reference markers are normally spaced at distances of 1km apart. According to Deacon et al. (1975) a spot length of 300m is adequate to accommodate reporting errors if reference markers are placed every kilometre and enforcement personnel are well trained.
- c) The spot length should be at least as long as the area of influence of a road hazard.

According to the South African Road Safety Manual (1999) an 'intersection' is defined as the intersection itself and the influence areas of the respective intersection approaches. The influence area of an approach is calculated by determining the safe stopping distance using the 85th percentile speed. Proposed influence areas of intersection for different speeds are shown in Table 5.1. (SA Road Safety Manual ; 1999).

Table 5.1 : Proposed influence areas of intersection

Speed	60 km/h	70 km/h	80 km/h	90 km/h	100 km/h
Influence length	87 m	115 m	147 m	185 m	224m

- d) The reliability in identifying hazardous locations is directly related to the spot length. As spot length increases, the probability of identifying a true hazardous location as hazardous increases and the probability of identifying a safe location as hazardous decreases.

- e) If spots are small it becomes increasingly difficult to make meaningful comparisons between spots because of the low accident counts.

On the basis of these considerations it is desirable that spot lengths be as long as possible. However if the spot is too long it could become excessively difficult and time-consuming to locate the hazard/s.

Deacon et al. (1975) recommends that a spot should preferably be between 480m and 800m long.

5.3.2 SECTIONS/SEGMENTS/ROUTES

It is important that *sections* be defined in a manner that ensures that the pavement condition, geometric design, traffic volume and other geometric, traffic and environmental variables etc. are uniform across the whole length of the section. *Sections* can be defined as a segment of road between two nodes, where nodes could be intersections or points where the pavement, geometric, traffic characteristics etc. change. According to this definition sections would normally be of variable length.

According to Deacon et al. (1975) observed accident rates are dependent on section length. High accident rates have been observed on short sections and low accident rates on long sections of roads. Long sections tend to have lower traffic volumes and fewer factors of traffic interference such as intersections, access points and changes in the number of lanes.

The variance of a Bayesian estimate depends amongst others on the exposure at a site. The longer the length of a section the larger the exposure and the smaller the variance. A smaller variance makes it easier to detect whether the estimate exceeds a threshold value.

EXAMPLE 5.1

Assume that the system wide average rate R_a from a group of similar sites is 1.1 acc/mvkm. Also assume that the harmonic mean (E^*) of all the exposures is 4.1 mvkm. Assume we have a section of road carrying 4000 vehicles per day. In order to illustrate the effect of segment length, $P(m_i > R_a)$ has been determined for segment lengths ranging from 1 km to 25 km. This procedure was also applied for different accident rates – 1.2, 1.3, 1.4, 1.6, 1.8 and 2 acc/mvkm.

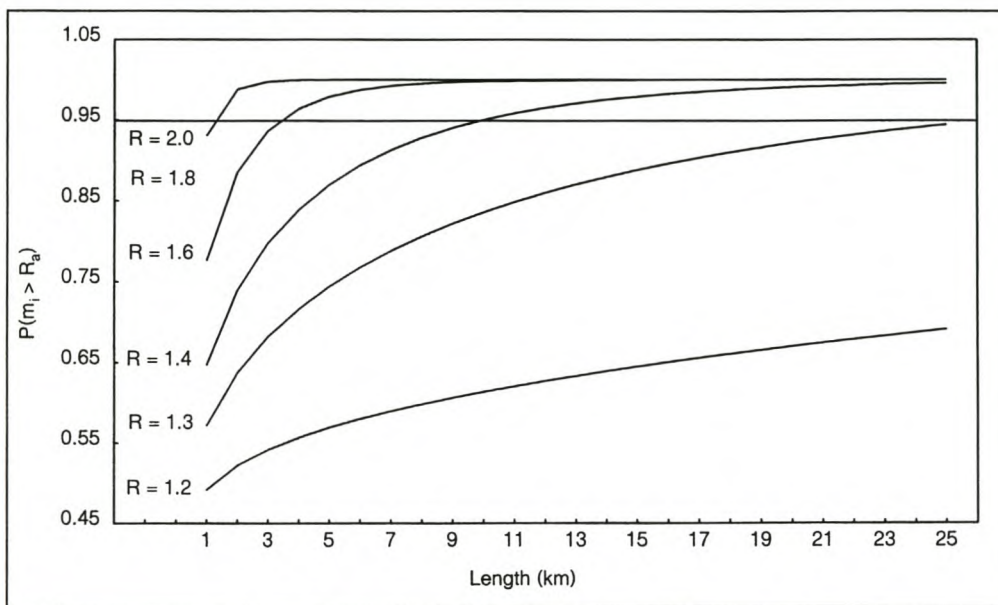


Figure 5.1 : The effect of segment length on $P(m_i > R_a)$.

It is evident that the higher the accident rate the shorter is the minimum length required for a site to be identified as a hazardous location.

For the reasons highlighted in Example 5.1 it is recommended that section lengths be constant. From the example it is evident that the shorter the length the higher the possible number of *false negative* identifications. With long section lengths the higher the possibility of an increase in the degree of *false positive* identifications.

5.4 IDENTIFICATION OF HAZARDOUS LOCATIONS

Hauer and Persaud (1984) have identified the following 4 criteria to evaluate the *efficiency* of a hazardous location identification method :

- The number of locations selected as hazardous.
- The number of truly hazardous sites selected for further investigation – *true positives*.
- The number of non-hazardous sites selected for further investigation – *false positives*.
- The number of truly hazardous sites NOT selected – *false negatives*.

The following Venn diagram graphically illustrates the above concepts.

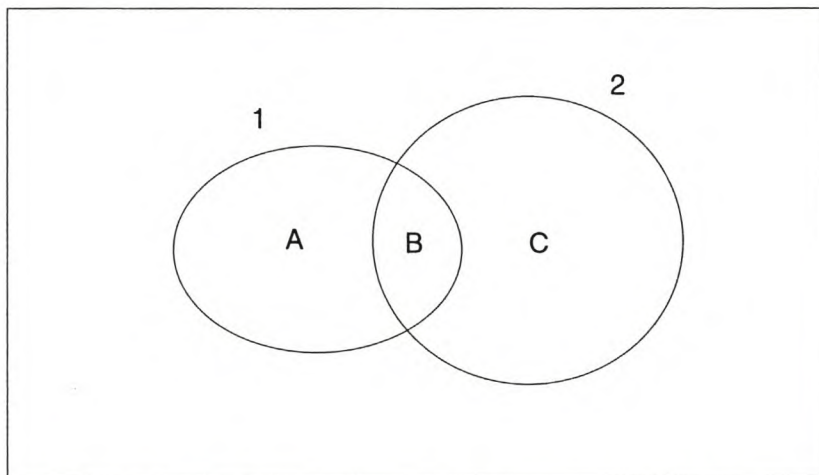


Figure 5.2 : A – False negatives , B – True positives and C – False positives

The total number of selected sites corresponds with the size of Curve 2. The true number of hazardous locations corresponds with the size of Curve 1. With a truly *efficient* identification method Curves 1 and Curve 2 will coincide.

An *efficient* 'sieve' , and therefore a good identification method, will maximise the number of *true positives* and minimise the number of *false positives* and *false negatives* (Hauer and Persaud ; 1984).

It is generally agreed (Higle and Hecht ; 1989) that a *false negative* error is far more serious than a *false positive* error. A *false negative* error would cause a true hazardous location to go untreated with potentially very severe consequences. A *false positive* error on the other hand would cause potentially scarce resources to be wasted on investigating a site that is not really a hazardous location, thereby reducing the amount of resources available to investigate true hazardous locations. According to Higle and Hecht (1989) a technique that tends to yield a low number of *false negatives* is good, as long as its is not accompanied by an excessively large number of *false positives*.

The different methods for identifying hazardous locations and for the preliminary ranking of sites locations can be grouped into two categories : *Conventional* methods and *Bayesian* methods.

5.4.1 CONVENTIONAL IDENTIFICATION METHODS

Palkowski and Menezes (1991) recommend using the following methods to identify hazardous locations :

- a) Accident Number method,
- b) Accident Severity method,
- c) Accident Rate method,
- d) Accident Rate-Number method,
- e) Quality Control method.

These methods are called *conventional methods* because they are based on the following 'conventional' assumptions :

- a) The observed accident rate or number is a reliable and valid measure of the true level of safety at a location i.e. $m_i = X_i$ or $m_i = R_i$.
- b) All sites within a reference group of similar sites have the same 'true' level of safety.

- c) The accident rate at a location is a constant irrespective of the level of exposure at the location i.e. a linear Safety Performance Function.

5.4.1.1 THE ACCIDENT NUMBER METHOD

This method considers only the frequency of accidents at a location. A location is classified as a hazardous location when the accident frequency exceeds a predetermined threshold.

This method is considered suitable for small town street systems, local street systems in larger cities and low volume rural roads. Because of the low and narrow range of traffic volumes on these types of roads the inclusion of the exposure factor is not critical (Roy Jorgensen Associates; 1975).

The primary purpose of the threshold value is to control the number of selected hazardous locations. The threshold value could be different from authority to authority depending on policy and available resources.

An increase in the threshold value will cause the number of incorrectly identified sites as well as the number of correctly identified sites to decrease. It appears that there is a trade off between the two and that a suitable threshold value will depend on the ability of a road authority to investigate all identified locations.

EXAMPLE 5.2

The experimental data in Appendix A1 was used to illustrate the impact of the chosen threshold value

The true levels of safety (m_t) for each of the 1000 locations were randomly generated from a Gamma distribution with an annual mean accident frequency of 4. All those locations where $P(m_t > 4) > 0.95$ are assumed to be true hazardous locations.

Using the randomly generated true level of safety m_t as the Poisson mean, 5 years of accident frequencies were randomly generated at each of the 1000 sites, for each of 5 study years.

For each study period the accident number (X) was determined. If the accident number exceeded a certain threshold value that location was then flagged as a hazardous location.

The next step was to, for each study period, determine the number of *true positives*, *false positives*, *false negatives* and *true negatives* for threshold values (T) of 5, 7 and 9 accidents per year. The following criteria were used :

True positive :	$P(m_t > 4) > 0.95$ AND $X \geq T$
False negative :	$P(m_t > 4) > 0.95$ AND $X < T$
True negative :	$P(m_t \leq 4) > 0.95$ AND $X < T$
False positive :	$P(m_t \leq 4) > 0.95$ AND $X \geq T$

As a measure of efficiency 3 different indicators were calculated. These are the i) number of sites identified as hazardous locations, ii) the *true positive* identification rate, and the iii) the *false positive* rate. The *true positive* rate is the % of hazardous locations correctly identified as hazardous while the *false positive* rate indicates the % of non-hazardous sites identified as hazardous .

Continue

Example 5.2 (continued)

Table 5.2 : Accident no. method : Efficiency assessment of different threshold values.

Collection period	1 Year	2 Year	3 Year	4 Year	5 Year
Threshold = 5 Accidents / year					
False negative	10	8	5	2	0
False positive	345	267	240	226	206
True negative	601	679	706	720	740
True positive	44	46	49	52	54
No. identified	389	313	289	278	260
% Incorrectly identified	36.5	28.2	25.4	23.9	21.8
% Correctly identified	81.5	85.2	90.7	96.3	100.0
Threshold = 7 Accidents / year					
False negative	27	25	23	23	20
False positive	130	70	50	38	30
True negative	816	876	896	908	916
True positive	27	29	31	31	34
No. identified	157	99	81	69	64
% incorrectly identified	13.7	7.4	5.3	4.0	3.2
% correctly identified	50.0	53.7	57.4	57.4	63.0
Threshold = 9 Accidents / year					
False negative	41	42	46	47	48
False positive	34	13	5	6	3
True negative	912	933	941	940	943
True positive	13	12	8	7	6
No. identified	47	25	13	13	9
% incorrectly identified	3.6	1.4	0.5	0.6	0.3
% correctly identified	24.1	22.2	14.8	13.0	11.1

From Table 5.2 it is evident that for the experimental data a low threshold value (5 accident / year) will yield a high % *correctly identified* as well as a high % *incorrectly identified*. Since in practice it is not known which sites were correctly or incorrectly identified all identified sites will have to be further investigated. If sufficient resources are available to investigate all identified sites then at least a high number of true hazardous locations will be subjected to an investigation – along with all the non-hazardous locations.

If a suitable reference group of sites is available a threshold value can be calculated from Equation 5.1.

METHOD CN1

$$X_T = \bar{X} + k\sigma_x \quad \dots[5.1]$$

where

\bar{X} - The sample mean accident frequency.

σ_x - Sample standard deviation.

k - A probability factor determined by the desired level of significance.

X_T - Threshold value.

A site is flagged as a hazardous location when $X_i > X_T$

The values of k corresponding to various levels of probability are shown in Table 5.3.

Table 5.3 : k-values

P	0.001	0.005	0.0075	0.05	0.075	0.10
K	3.09	2.576	1.96	1.645	1.440	1.282

Source : Palkowski and Menezes ; 1991

A probability level is selected to ensure that an accident number/rate is sufficiently large so that it cannot be reasonably attributed to random occurrences (Palkowski and Menezes ; 1991). The probability level serves the same function as a *threshold* value.

Method CN1 is based on the assumption that accident frequencies between sites are distributed according to the Normal distribution. This assumption, according to Hagle and Hecht (1989), causes this method to consistently yield a high number of *false negative* identifications.

Since the accident frequency at a site is a function of the traffic volume, the accident number method tends to favour sites with high traffic volumes.

Two locations with the same amount of accidents, but with different traffic volumes obviously have different levels of risk. The *accident rate* method considers this variable.

5.4.1.2 THE ACCIDENT SEVERITY METHOD

This method is a variant of the *accident number* method, the difference being that a site is considered a hazardous location if the Equivalent Accident Number (EAN) exceeds a predetermined threshold level, where :

$$EAN = w_f(F) + w_{sr}(SR) + w_{sl}(SL) + D \quad \dots[5.2]$$

Where

w_f, w_{sr}, w_{sl} - *The weights associated with fatal, serious injury and slight injury accidents respectively.*

F, SR, SL, D - *The number of fatal, serious injury, slight injury and damage only accidents respectively.*

The weighting factors represent the average cost of the respective accident severity classes relative to the cost of a *Damage Only* accident.

The K21 Manual (Opperman and Hutton ; 1991) proposes the following weighting factors to be used for the calculation of the EAN :

Fatal accidents	12
Injury accidents	3
Damage only accidents	1

The average 1998 cost of South African road traffic accidents, according to Schutte (2000), are shown in Table 5.4.

Table 5.4 : Cost of accidents in 1998.

Severity of accident	Average cost per collision	Weighting factor
Fatal	R 388 487	25
Serious Injury	R 88 248	6
Slight Injury	R 23 723	1.5
Damage Only	R 15 694	1

The total cost of accidents at a location can be obtained by multiplying the EAN by the cost of a *Damage Only* accident.

EXAMPLE 5.3

Consider the following two locations :

Table 5.5 : Accident data

Location	Fatal	Serious	Slight	Damage	Total
A	2	22	25	45	94
B	5	15	18	40	78

The EAN for these sites using the K21 manual weighting factors are:

Location A 210
Location B 199

Using the *weighting factors* derived from the accident costs the EAN for the two sites are as follows :

Location A 265 (Cost = 265 * 15694 = R 4 158 910)
Location B 282 (Cost = 282 * 15694 = R 4 425 708)

Using only the *accident number* method Location A will be considered worse than Location B. Using the 'K21' factors indicates that Location A is worse than Location B, while using the 'Cost' factors would indicate that Location B is the worst location.

The choice of appropriate *weighting factors* is critical to this method.

The EAN method recognises accident *severity* as a prime factor in the identification of hazardous locations.

According to Deacon et al. (1975) the EAN represents not only the accident number but also the severity of these accidents. The EAN method therefore favours locations with high accident numbers (i.e. high exposure sites) with a relatively large proportion of serious and fatal accidents. Deacon et al. (1975) states that locations selected using the EAN method will be more economically efficient than sites selected using the *accident number* method or the *accident rate* method.

5.4.1.3 THE ACCIDENT RATE METHOD

This method is very similar to the *accident number* method. Instead of using accident frequencies, accident rates are used. A location is identified as hazardous when its accident rate exceeds some predetermined threshold.

This threshold (R_T) can be determined as follows (Highle and Hecht ; 1989):

METHOD CR1

$$R_T = \bar{R} + k\sigma_R \quad \dots[5.4]$$

where

\bar{R} - Sample mean of accident rates of all reference group sites.

σ_R - Sample standard deviation of accident rates.

k - Probability factor.

Location i is flagged as a hazardous location if $R_i > R_T$.

Using this method requires sufficient information to reliably determine the *exposure* at each location in the network.

According to McGuigan (1982) the *accident rate* method tends to produce a set of locations biased towards low accident totals and low traffic usage. By choosing an appropriate threshold level it is possible to limit the number of selected sites with low accident numbers. The choice of an appropriate threshold can be problematic in that a too low level could cause locations with limited accident reduction potential to be included or too high a level could cause locations with good accident reduction potential to be excluded. The identification and ranking of hazardous locations using accident rates therefore requires a high degree of subjective decision making by the analyst.

5.4.1.4 NUMBER-RATE METHOD

In the preceding section it was stated that the *accident number* method is biased towards high volume sites and the *accident rate* method towards low volume sites.

The *number-rate (NR) method* attempts to find a compromise between the *accident number* method and the *accident rate* method.

The *NR method* is based on the concept that those sites that are flagged as hazardous locations using the *accident number* method (e.g. Method CN1) **AND** the *accident rate* method (e.g. Method CR1) can be considered abnormal and therefore truly hazardous.

The number of sites so identified can be manipulated by adjusting the threshold values. According to Barbaresso et al.(1982) sites identified by the *NR method* can be ranked according to their *Severity Indices* (SI's), where ;

$$SI = \frac{EAN}{N} \quad \dots[5.5]$$

SI - *Severity Index*.

EAN - *Equivalent Accident Number* ; N - *Total number of accidents*.

EXAMPLE 5.4

The sample means, standard deviations and threshold values (assuming $k = 1.645$) for accident numbers (acc/km) and accident rates (acc/mvkm) of all 113 segments in Appendix B1 and B2 are shown below :

Table 5.6 : Mean, standard deviation and threshold values

Safety Measure	Mean	St. Dev.	Threshold
Accident number	5.6	5.4	14.5
Accident rate	1.15	1.1	3.0

Table 5.7 indicates all those sites where the accident number threshold OR the accident frequency threshold are exceeded.

Table 5.7 : List of hazardous locations from Methods CN1, CR1 and NR

Road	Start	End	L	Acc/km	R	CN1	CR1	NR
NR00205	51.88	52.62	0.74	31.08	7.87	1	1	1
MR00027	51.73	52.29	0.56	26.79	4.34	1	1	1
MR00165	0	3.63	3.63	8.82	4.02	0	1	0
MR00227	5.89	9.66	3.77	5.57	3.43	0	1	0
TR02801	0	2.14	2.14	5.61	3.31	0	1	0
TR03201	44.35	45.15	0.8	8.75	3.40	0	1	0
MR00027	51.15	51.73	0.58	29.31	2.83	1	0	0
MR00027	67.19	68.56	1.37	18.25	2.65	1	0	0
MR00165	3.63	7.47	3.84	19.79	2.45	1	0	0
NR00205	9.85	11.4	1.55	18.71	2.29	1	0	0
TR00202	37.09	42.36	5.27	15.75	1.48	1	0	0

It is evident that the RN method only identified two sites as hazardous.

Since both the *CN1* and *CR1* methods are associated with a high degree of *false negative* identifications, and the *RN method* tends to identify only those sites common to methods *CN1* and *CR1*, it can be expected that the degree of *false negative* identifications using the *RN method* could be considerably worse than that of the *CN1* and *CR1* methods.

5.4.1.5 THE QUALITY CONTROL METHOD

The *quality control* method calculates, using statistical techniques, a critical accident rate for each site under consideration. The value of the critical rate is a function of the system wide average accident number/rate for all reference groups, the exposure at a site and a probability factor k .

a) THE NUMBER QUALITY CONTROL METHOD

METHOD CN2

$$X_{cr} = X_a + k \cdot \sqrt{X_a} + 0.5 \quad \dots[5.6]$$

X_{cr} = Critical accident number.

X_a = Mean (average) accident number for all reference locations

k = Probability factor.

Location i is flagged as a hazardous location if $X_i > X_{cr}$.

Method CN2 is based on the assumption that the distribution of accident numbers between sites follows a Poisson distribution and that this distribution can be estimated, using *the Central Limit Theorem*, by a Normal distribution with mean and variance equal to the system wide average X_a .

To illustrate the *efficiency* of this method the experimental data has been used to estimate the number of *false negatives*, *false positives*, *true negatives* and *true positives* using Equation 5.6 for different study periods ranging from 1 to 5 years. See Example 5.5 .

EXAMPLE 5.5

The threshold value (X_{cr}) for each study period was estimated from Eqn. 5.6 with $k = 1.645$.

Table 5.8 : Method CN2 - threshold values

Collection Period	X_a	X_{cr}
1 Year	4	7.9
2 Year	8	13.2
3 Year	12	18.2
4 Year	16.1	23.2
5 Year	20.2	28.1

For a location to be flagged as a hazardous location $X_i > X_{cr}$.

The next step was to, for each collection period, determine the number of *true positives*, *false positives*, *false negatives* and *true negatives*. The following criteria were used :

- True positive : $P(m_t > 4) > 0.95$ AND $X > X_{cr}$
- False negative : $P(m_t > 4) > 0.95$ AND $X < X_{cr}$
- True negative : $P(m_t \leq 4) > 0.95$ AND $X < X_{cr}$
- False positive : $P(m_t \leq 4) > 0.95$ AND $X > X_{cr}$

Table 5.9 : Method CN2 – efficiency assessment

Collection period	1 Year	2 Year	3 Year	4 Year	5 Year
False negatives	37	25	19	9	3
False positives	75	70	82	100	114
True negative	871	876	864	846	832
True positives	17	29	35	45	51
No. identified	92	99	117	145	165
% Incorrectly identified	7.9	7.4	8.7	10.6	12.1
% Correctly Identified	31.5	53.7	64.8	83.3	94.4

It is evident that as the collection period increases the number of *true positive* identifications tends to increase. This however is achieved at the expense of an increased number of *false positive* identifications.

An alternative to using *Method CN2* is *Method CP1*. *Method CP1* is based on the assumption that accident numbers between sites are Poisson distributed around the system wide average X_a .

METHOD CP1

$$P(X > X_a) = \sum_{i=0}^X \frac{X_a^i e^{-X_a}}{i!} \quad \dots[5.7]$$

A site, which experienced X_i accidents, is considered to be a hazardous location when $P(X_i > X_a) > 0.95$. (Assuming a 95 % degree of confidence.)

With modern computer technology and the use of spreadsheets $P(X > X_a)$ can be calculated directly using Equation 5.7 without resorting to the *Central Limit Theorem* and *Method CN2*. Generally, because of the difficulty of solving Equation 5.7 this equation can be approximated by Equation 5.6.

EXAMPLE 5.6

Assuming $X_a = 6$ and $X = 11$.

Then $P(X > X_a)$ can be determined using the following Microsoft® Excel® function :

=POISSON(11,6,TRUE) = 0.98.

To illustrate the *efficiency* of this *Method CP1* the experimental data has been used to estimate the number of *false negatives*, *false positives*, *true negatives* and *true positives* using Equation 5.7 for different study periods ranging from 1 to 5 years. See Example 5.7 and Appendix A2.

EXAMPLE 5.7

The average number (X_a) and the threshold value (X_{cr}) for each study period are indicated in Table 5.10. X_{cr} was determined as the value of X_i in Equation 5.7 for which $P(X_i > X_a)$.

Table 5.10 : Method CP1 – Critical values

Collection Period	X_a	X_{cr}
1 Year	4	8
2 Year	8	13
3 Year	12	18
4 Year	16.1	23
5 Year	20.2	28

For a location to be flagged as a hazardous location : $X \geq X_{cr}$.

Table 5.11 : Method CP1 – efficiency assessment

Collection Period	1 Year	2 Year	3 Year	4 Year	5 Year
False negative	37	23	12	8	2
False positive	75	99	117	119	138
True negative	871	847	829	827	808
True positive	17	31	42	46	52
No. identified	92	130	159	165	190
% Incorrectly identified	7.93	10.47	12.37	12.58	14.59
% Correctly identified	31.5	57.4	77.8	85.2	96.3

Comparing the results of Example 5.6 with the results of this Example it is evident that *Method CP1* identifies less sites than method CN2. Take for example Year 3 – *Method CP1* will identify all those sites with 18 accidents or more, while method CN2 will only identify those sites with 19 accidents or more. Since only some of the sites that had 18 accidents are true hazardous locations, method CP1 will identify more of the truly hazardous locations than method CN2 but also more of the non-hazardous locations.

b) THE RATE QUALITY CONTROL METHOD

METHOD CR2

$$R_{cr}(i) = R_a + k \sqrt{\frac{R_a}{E(i)}} + 0.5/E(i) \quad \dots[5.8]$$

$R_{cr}(i) =$ Critical accident rate for site i .

$R_a =$ Average accident rate of reference group

$E(i) =$ Exposure for site i

$$R_a = \frac{\sum X(i)}{\sum E(i)} \quad \dots[5.9]$$

A site is identified as a hazardous location if $R(i) \geq R_{cr}(i)$

An alternative to Method CR2 is the direct Method CP2.

METHOD CP2

$$P(R_i > R_a) = \sum_{i=0}^{X_i} \frac{(R_a E_i)^{X_i} e^{-(R_a E_i)}}{i!} \quad \dots[5.10]$$

Method CP2 is based on the assumption that the number of accidents at a location is Poisson distributed around the mean which is equal to $R_a * E_i$.

The number and rate quality control methods assume that the observed accident number at a location has a Poisson distribution about the system wide average (X_a or $R_a * E_i$), and that this system wide average accurately represents the level of safety at a location. This assumption is not correct, since according to Abbess et al. (1981)

the 'true' level of safety at a site between the members of a reference group generally follow a Gamma distribution around the system wide average, and that the observed accident number at a location is Poisson distributed around the 'true' level of safety at a site and not around the system wide average. Accident frequencies therefore according to Vogt and Bared (1998) follow a Negative Binomial distribution between the sites of a reference group. Whereas with the Poisson distribution the variance is equal to the mean, with the Negative Binomial distribution the variance is larger than the mean. The *quality control* methods therefore use a variance that is underestimated, with the consequence that the critical (threshold) values are also underestimated. Although this could cause a high degree of *true positive* identification it could also cause a high degree of *false positive* identification.

EXAMPLE 5.8

Consider the data in Appendix B1 and B2. The total number of accidents on the 113 segments over a 4 year period = 2433 and the total exposure on all these segments = 3270.3 million-vehicle-kilometres.

$$R_a = 2433/3270.3 = 0.74 \text{ accidents/mvkm}$$

Using the above value of R_a and the level of exposure at each site (E_i) the critical rate for each segment was determined using Equation 5.8.

Table 5.12 : Method CR2 : Top 10 segments where $R > R_{cr}$

Road	Start	End	AADT	E	X	R	R_{cr}
NR00205	51.88	52.62	2704	2.92	23	7.87	1.74
MR00027	51.73	52.29	4220	3.45	15	4.34	1.65
MR00165	0	3.63	1502	7.97	32	4.02	1.31
MR00227	5.89	9.66	1111	6.12	21	3.43	1.40
TR03201	44.35	45.15	1760	2.06	7	3.40	1.98
TR02801	0	2.14	1158	3.62	12	3.31	1.63
MR00027	51.15	51.73	7077	6.00	17	2.83	1.41
MR00223	6.3	9.63	1618	7.87	21	2.67	1.31
TR03201	42.84	44.35	1708	3.77	10	2.65	1.61
MR00027	67.19	68.56	4722	9.45	25	2.65	1.26
NR00108	2.68	3.61	2530	3.44	9	2.62	1.65

5.4.2 BAYESIAN IDENTIFICATION METHODS

According to Persaud et al. (1999b) the conventional methods listed above are known to have difficulties in identifying hazardous locations because of the potential bias due to the *regression-to-mean* phenomenon in which sites with a randomly high accident count can be wrongly identified as being hazardous and vice versa.

According to Abbess et al. (1981) the *regression-to-mean* effect (random error) has been eliminated from road safety estimates that have been obtained from using Empirical Bayesian methods. The difficulties associated with the conventional methods as mentioned by Persaud et al. (1999b) can be overcome by using Bayesian safety estimates.

Higle and Hecht (1989) investigated and evaluated 4 different Bayesian methods to identify hazardous locations :

i) METHOD B1

According Higle and Witkowski (1988) location i is hazardous if the probability is greater than δ that its true accident rate/number m_i , exceeds the observed average rate/number (\bar{Z}) across the reference population. (Z refer to both R and X)

$$P(m_i > \bar{Z}) > \delta \quad \dots[5.11]$$

where

m_i - Bayesian estimate of safety.

δ - The desired level of confidence e.g. 0.90, 0.95, 0.99

\bar{Z} - The sample mean level of all Z_i 's

$$\bar{Z} = \frac{1}{n} \sum_{i=1}^n Z_i \quad \dots[5.12]$$

Since the Bayesian estimate of safety, m_i , has a Gamma distribution :

$$P(m_i > \bar{Z}) = 1 - P(m_i \leq \bar{Z}) = 1 - \int_0^{\bar{Z}} \frac{\beta_i^{\alpha_i} m^{\alpha_i-1} e^{-\beta_i m}}{\Gamma(\alpha_i)} dm \quad \dots[5.13]$$

EXAMPLE 5.9

The experimental data in Appendix A1 was used to illustrate the efficiency of *Method B1* in identifying hazardous locations for different study periods. See Appendix A3.

The true levels of safety (m_i) for each of the 1000 locations were randomly generated from a Gamma distribution with an annual mean accident frequency of 4. All those locations where $P(m_i > 4)$ are assumed to be true hazardous locations.

Using the randomly generated true level of safety m_i as the Poisson mean 5 years of accident frequencies were randomly generated at each of the 1000 sites. For each study period ranging from 1 to 5 years, Bayesian estimates (m_i) and their gamma parameters (α, β) were obtained for each site using the *method of sample moments*. The mean accident frequency \bar{X} was also determined for each collection period. For each site $P(m_i > \bar{X})$ was determined using the GAMMADIST function of Microsoft® Excel® . A site was flagged as hazardous if $P(m > \bar{X}) > 0.95$.

The following criteria were used to determine the number of true *positive*, *false negative*, *true positive* and *false positive* identifications respectively.

TRUE POSITIVE :	$P(m_i > 4) > 0.95$ AND $P(m_i > \bar{X}) > 0.95$
FALSE NEGATIVE :	$P(m_i > 4) > 0.95$ AND $P(m_i > \bar{X}) \leq 0.95$
TRUE NEGATIVE :	$P(m_i > 4) \leq 0.95$ AND $P(m_i > \bar{X}) \leq 0.95$
FALSE POSITIVE :	$P(m_i > 4) \leq 0.95$ AND $P(m_i > \bar{X}) > 0.95$

Continue ...

Example 5.9 (continued)

As a measure of efficiency 3 different indicators were calculated. These are the i) number of sites identified as hazardous locations, ii) the *true positive* identification rate, and the iii) the *false positive* rate. The true positive rate is the % of hazardous locations correctly identified as hazardous while the *false positive* rate indicate the % of non-hazardous locations identified as hazardous.

Table 5.13 : Method B1 – efficiency assessment

Identification	1 Year	2 Year	3 Year	4 Year	5 Year
True positive	3	19	31	40	45
False negative	51	35	23	14	9
True negative	938	915	896	884	876
False positive	8	31	50	62	70
Total identified	11	50	81	102	115
% Correctly identified	5.6	35.2	57.4	74.1	83.3
% Incorrectly identified	0.85	3.28	5.29	6.55	7.40

ii) Method B2

According to Hagle and Witkowski (1988) location *i* is hazardous if the probability is greater than δ that its true accident rate m_i , exceeds the observed regional accident rate R_a .

$$P(m_i > R_a) > \delta \tag{3.14}$$

$$R_a = \frac{\sum_{i=1}^n X_i}{\sum_{i=1}^n E_i} \tag{3.15}$$

Since the Bayesian estimate of safety, m_i , has a Gamma distribution :

$$P(m_i > R_a) = 1 - P(m_i \leq R_a) = 1 - \int_0^{R_a} \frac{\beta_i^{\alpha_i} m^{\alpha_i-1} e^{-\beta_i m}}{\Gamma(\alpha_i)} dm \tag{3.16}$$

EXAMPLE 5.10

Consider the data in Appendix B1 and B2.

The mean accident rate for the 113 segments is $\bar{R} = 1.15$ acc/mvkm and the area wide average rate $R_a = 0.74$ acc/mvkm. The reason for such a large difference between \bar{R} and R_a lies in the fact that in the sample of sites used for the estimation of these rates there is a wide variation in the magnitude of exposures. This is particularly the case if the road segments in the sample have different lengths.

Table 5.14 shows the $P(m > R_a)$ and $P(m > \bar{R})$ for the top ten sites on the basis of their Bayesian estimates, m_i .

Table 5.14 : Method B2 – top 10 hazardous locations

Road	Start	End	E	R	m	α'	β'	$P(m > \bar{R})$	$P(m > R_a)$
NR00205	51.88	52.62	2.92	7.87	5.98	4.06	24.31	1.000	1.000
MR00165	0	3.63	7.97	4.02	3.66	9.11	33.31	1.000	1.000
MR00027	51.73	52.29	3.45	4.34	3.55	4.59	16.31	1.000	1.000
MR00227	5.89	9.66	6.12	3.43	3.07	7.26	22.31	1.000	1.000
TR02801	0	2.14	3.62	3.31	2.80	4.76	13.31	0.997	1.000
TR03201	44.35	45.15	2.06	3.40	2.60	3.20	8.31	0.974	0.998
MR00027	51.15	51.73	6.00	2.83	2.57	7.14	18.31	0.998	1.000
MR00027	67.19	68.56	9.45	2.65	2.48	10.59	26.31	1.000	1.000
MR00223	6.3	9.63	7.87	2.67	2.48	9.01	22.31	0.999	1.000
MR00165	3.63	7.47	31.02	2.45	2.40	32.17	77.31	1.000	1.000

$P(m > \bar{R})$ was determined using the following expression in Microsoft® Excel® :

=1-GAMMADIST(1.15, β , 1/ α , TRUE)

And $P(m > R_a)$ was determined using :

=1- GAMMADIST(0.74, β , 1/ α , TRUE)

5.4.3 COMPARISON OF IDENTIFICATION METHODS

EXAMPLE 5.11

The *true positive* rates of the CP1, CN2 and B1 methods as determined in Examples 5.5, 5.7 and 5.9 compared in Figure 5.3 The false positive rates of the P1, CN2 and B1 methods are compared in Figure 5.4.

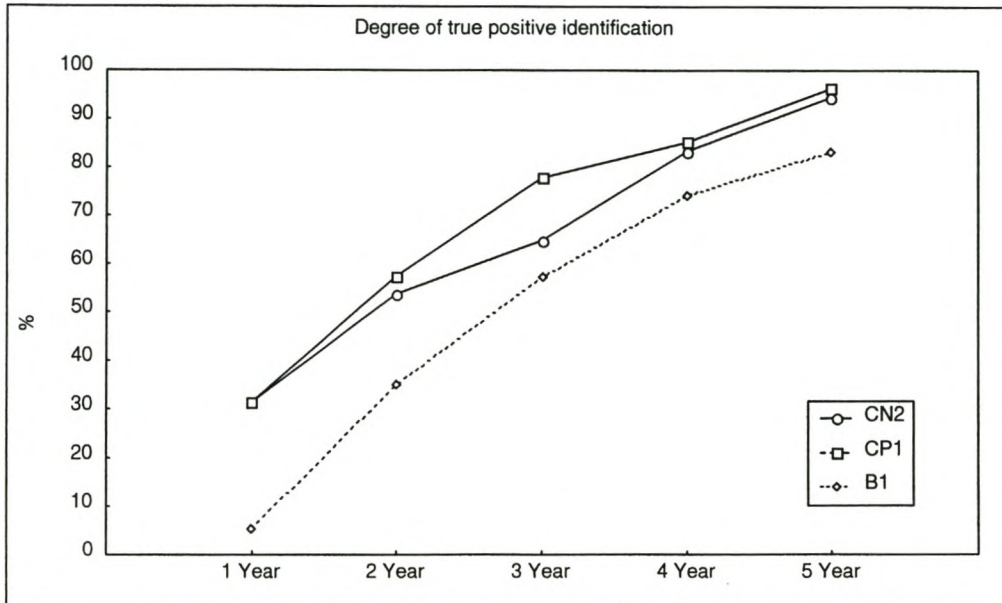


Figure 5.3 : Degree of *true positive* identifications

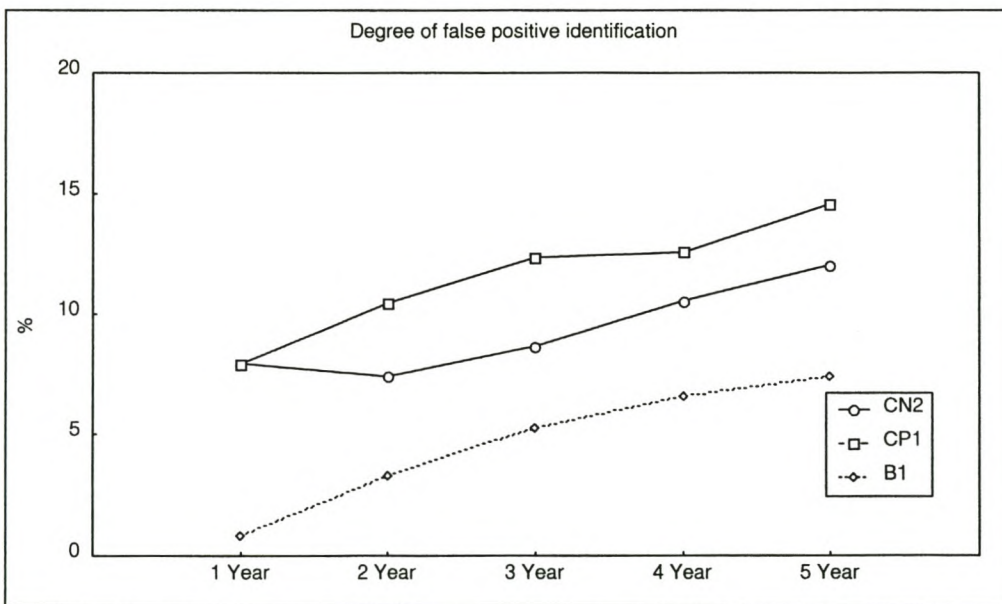


Figure 5.4 : Degree of *false positive* identifications.

Continue

EXAMPLE 5.11 (Continued)

From Figures 5.3 and 5.4 it is evident that for at least the experimental data that method B1 generally has a lower true positive rate than CP1 and CN2, but that this is somewhat compensated for by a lower rate of *false positive* identifications. This means that although method B1 identified less true hazardous locations, it also identified significantly less sites which are not true hazardous locations.

Although methods CP1 and CN2 have similar rates of *true positive* identifications, method CN2 performed better with regards to *false positive* identifications.

Higle and Witkowski(1988) and Higle and Hecht (1989) conducted a detailed experimental study to compare the efficiency of the CR1, CR2, B1 and B2 identification methods.

They came to the following conclusions :

- a) Methods CR2, B1 and B2 correctly identify a significantly higher fraction of the truly hazardous sites than does CR1.
- b) Because of the underlying assumption of normality in the distribution of accident rates/numbers method CR1 tends to identify a larger number of *false negatives* than methods CR2, B1 and B2.
- c) The CR2, B1 and B2 methods tend to be equally efficient. Each yield low numbers of *false negative* identifications and are generally successful in selecting sites that are truly hazardous. This however comes at the expense of an increase in the number of *false positive* identifications, which may be a result of a sensitivity to the traffic volume at a site.
- d) Method B2 is more efficient than CR2 when data are sparse or when numerous years of comparable data are not available.

EXAMPLE 5.12

The CR1, CR2, B1 and B2 methods of hazardous location identification were applied to the sample of 113 road segments in Appendix B1 and B2. All the locations that were identified as a hazardous location of one or more of these methods are listed in Table 5.15

Method CR1 flagged only 6 of the 113 sites as hazardous. This seems to indicate the likelihood of a high rate of *false negative* identifications, as was also concluded by Higle and Hecht (1989) as well as Higle and Witkowski (1988).

Method CR2 identified a total of 48 sites as opposed to the 34 by Method B2. It is likely that method CR2 identified most of the true hazardous locations but also a large number of non-hazardous locations, while method B2 might have identified less hazardous locations, but also less non-hazardous locations.

It appears as if Method B2 is sensitive to the level of exposure at a site. When combining the *prior* estimates with the accident frequency and exposure level using *Bayes theorem* the larger the level of exposure the smaller the variance of the Bayesian estimate m_i . The smaller the variance, the smaller is the critical threshold value. To illustrate the following two segments can be compared :

Road	Start	End	E	m	VAR(m)	B2
NR00205	40.64	49.34	38.93	1.48	0.037	1
NR00205	58.6	59.21	3.14	1.48	0.345	0

Both these locations have the same Bayesian estimates = 1.48 acc/mvkm but only the first location was flagged as a hazardous location because of its relatively small variance (= 0.037), which in turn is a result of its high exposure level = 38.93 mvkm.

Method B2 therefore favour sites with a potentially high PAR (Potential Accident Reduction). It also tends to favour sites which are longer than other sites but not necessarily more unsafe. It is advisable to, when applying the B2 method, to define the locations in a manner which ensures that they are all of the same length.

Continue ...

Example 5.12 (continued)

Table 5.15 : Comparison of CR1, CR2, B1 and B2 identification methods

Road	Start	End	E	R	m	CR1	CR2	B1	B2
NR00205	51.88	52.62	2.92	7.87	5.98	1	1	1	1
MR00165	0	3.63	7.97	4.02	3.66	1	1	1	1
MR00027	51.73	52.29	3.45	4.34	3.55	1	1	1	1
MR00227	5.89	9.66	6.12	3.43	3.07	1	1	1	1
TR02801	0	2.14	3.62	3.31	2.80	1	1	1	1
TR03201	44.35	45.15	2.06	3.40	2.60	1	1	1	1
MR00027	51.15	51.73	6.00	2.83	2.57	0	1	1	1
MR00027	67.19	68.56	9.45	2.65	2.48	0	1	1	1
MR00223	6.3	9.63	7.87	2.67	2.48	0	1	1	1
MR00165	3.63	7.47	31.02	2.45	2.40	0	1	1	1
TR03201	42.84	44.35	3.77	2.65	2.31	0	1	1	1
NR00108	2.68	3.61	3.44	2.62	2.25	0	1	1	1
NR00205	9.85	11.4	12.67	2.29	2.20	0	1	1	1
TR03201	0	5.71	14.42	2.22	2.14	0	1	1	1
NR00205	49.34	51.8	14.76	2.17	2.10	0	1	1	1
TR02801	2.14	3.73	8.74	2.17	2.06	0	1	1	1
DR01056	0	1.32	2.14	2.34	1.93	0	1	0	1
TR00204	50.54	55.03	17.47	1.95	1.90	0	1	1	1
MR00313	1.6	3.1	4.04	1.98	1.80	0	1	0	1
TR00204	44.28	45.16	3.15	1.91	1.71	0	1	0	1
TR02801	16.37	17.18	4.56	1.76	1.64	0	1	0	1
NR00205	52.62	58.6	30.12	1.56	1.55	0	1	1	1
TR02801	17.18	19.93	16.19	1.54	1.52	0	1	0	1
NR00205	40.64	49.34	38.93	1.49	1.48	0	1	1	1
NR00205	58.6	59.21	3.14	1.59	1.48	0	1	0	0
TR00202	37.09	42.36	56.21	1.48	1.47	0	1	1	1
TR03302	2.74	5.01	4.53	1.55	1.47	0	1	0	0
TR02801	19.93	23.6	22.73	1.45	1.44	0	1	0	1
NR00205	0	2.02	12.44	1.45	1.42	0	1	0	1
DR01101	1.9	5.79	11.25	1.33	1.32	0	1	0	1
TR00204	2.84	4.4	6.79	1.32	1.30	0	1	0	0
NR00105	29.83	31.89	7.86	1.27	1.26	0	1	0	0
DR01105	0	3.99	39.16	1.25	1.25	0	1	0	1
TR02801	3.73	9.1	23.78	1.22	1.22	0	1	0	1
MR00216	0	3.45	22.78	1.19	1.18	0	1	0	1
TR00204	45.16	50.54	19.59	1.17	1.17	0	1	0	1
TR02801	23.6	25.83	15.65	1.15	1.15	0	1	0	1
DR01101	0	1.9	10.49	1.14	1.14	0	1	0	0
TR02901	39.92	45.21	8.08	1.11	1.12	0	1	0	0
NR00205	7.12	9.12	18.06	1.00	1.01	0	1	0	0
MR00187	6.96	8.03	9.21	0.98	1.00	0	1	0	0
NR00208	65.71	67.78	16.38	0.92	0.93	0	1	0	0
NR00107	0	34.86	128.65	0.91	0.91	0	1	0	1
TR03201	5.71	15.26	21.85	0.87	0.88	0	1	0	0
TR03201	15.26	22.79	15.09	0.86	0.88	0	1	0	0
MR00191	16.66	20.2	19.69	0.86	0.88	0	1	0	0
NR00108	62	72.44	37.60	0.85	0.86	0	1	0	0
TR03102	1.94	9.04	37.92	0.82	0.83	0	1	0	0
TOTAL						6	48	20	34

5.5 RANKING METHODOLOGIES

The ranking of hazardous locations is important when a road authority does not have sufficient resources to investigate and treat all identified hazardous locations. It could be advantageous to have two levels of ranking, a preliminary level and a final level of ranking. The objective of the preliminary level of ranking should be to determine which sites to investigate further, given the resource constraints. The investigation of the selected sites should yield an estimate of the expected economic return (e.g. Benefit/Cost ratio, Internal Rate of Return etc.) as a result of the selected remedial measures. The final level of ranking is then based on these economic indicators to determine which combination of sites, when treated, would provide the best economic returns. To prevent the unnecessary investigation of hazardous locations the ranking index used for the preliminary ranking procedure should provide a good measure of the anticipated economic return.

Some of the ranking methods that will be presented in this section can also be used as a way to identify locations to apply remedial measures to, without going through the formal identification procedures as presented in the preceding sections.

5.5.1 ACCIDENT RATE/NUMBER RANKING METHOD

According to McGuigan (1982) using the accident rate as a ranking criterion will produce a bias towards sites with low accident totals and low traffic flows. Using the accident number on the other hand will produce a bias towards sites with high traffic volumes and high accident numbers.

EXAMPLE 5.13

Consider the information in Tables 5.16 and Table 5.17. The road segments in Table 5.16 are ranked according to their accident numbers while the segments in Table 5.17 are ranked according to their accident rates. For each site the PAR index has also been determined (see Paragraph 5.5.3).

Table 5.16 : Top 10 sites : Accident Number method

Road	Start Km	End Km	AADT	N	N Rank	R	R Rank	E	PAR
NR2/5	51.88	52.62	2704	31.08	1	7.87	1	2.9	20.84
MR27	51.15	51.73	7077	29.31	2	2.83	7	6.0	12.56
MR27	51.73	52.29	4220	26.79	3	4.34	2	3.5	12.45
MR165	3.63	7.47	5530	19.79	4	2.45	12	31.0	53.04
NR2/5	9.85	11.4	5593	18.71	5	2.29	14	12.7	19.63
MR27	67.19	68.56	4722	18.25	6	2.65	10	9.5	18.01
TR2/2	37.09	42.36	7300	15.75	7	1.48	27	56.2	41.41
NR2/5	49.34	51.8	4107	13.01	8	2.17	17	14.8	21.08
DR1105	0	3.99	6717	12.28	9	1.25	33	39.2	20.02
TR28/1	2.14	3.73	3762	11.95	10	2.17	16	8.7	12.53
TOTALS			51732	196.91		29.50		184.4	231.56

Table 5.17: Top 10 sites : Accident Rate method

Road	Start Km	End Km	AADT	R	R Rank	N	R Rank	E	PAR
NR2/5	51.88	52.62	2704	7.87	1	31.08	1	2.92	20.84
MR27	51.73	52.29	4220	4.34	2	26.79	3	3.45	12.45
MR165	0	3.63	1502	4.02	3	8.82	17	7.97	26.11
MR227	5.89	9.66	1111	3.43	4	5.57	41	6.12	16.47
TR32/1	44.35	45.15	1760	3.40	5	8.75	18	2.06	5.48
TR28/1	0	2.14	1158	3.31	6	5.61	39	3.62	9.32
MR27	51.15	51.73	7077	2.83	7	29.31	2	6.00	12.56
MR223	6.3	9.63	1618	2.67	8	6.31	37	7.87	15.17
TR32/1	42.84	44.35	1708	2.65	9	6.62	34	3.77	7.21
MR27	67.19	68.56	4722	2.65	10	18.25	6	9.45	18.01
TOTALS			27580	37.18		147.1		53.23	143.61

The sum of the AADT's for the locations selected using the accident number method is considerably larger than those selected using the accident rate method, confirming that the accident rate method is biased towards low volume sites and the accident number method towards high volume sites. As a result the total number of observed accidents at the sites selected using the accident number method is also larger than that of the sites selected using the accident rate method.

5.5.2 BAYESIAN SAFETY ESTIMATE METHOD

This method ranks sites according to their Bayesian estimates (m).

EXAMPLE 5.14

In Appendix B1 the Bayesian safety estimates were determined for each of the 113 sites using the *method of sample moments*. Table 5.18 shows the top 10 sites ranked according to their Bayesian safety estimates.

Table 5.18 : Top 10 sites – Bayesian estimate method

Road	Start	End	E	AADT	R	R Rank	m	M Rank	w ₁
NR00205	51.88	52.62	2.92	2704	7.87	1	5.98	1	0.72
MR00027	51.73	52.29	3.45	4220	4.34	2	3.55	3	0.75
MR00165	0	3.63	7.97	1502	4.02	3	3.66	2	0.87
MR00227	5.89	9.66	6.12	1111	3.43	4	3.07	4	0.84
TR03201	44.35	45.15	2.06	1760	3.40	5	2.60	6	0.64
TR02801	0	2.14	3.62	1158	3.31	6	2.80	5	0.76
MR00027	51.15	51.73	6.00	7077	2.83	7	2.57	7	0.84
MR00223	6.3	9.63	7.87	1618	2.67	8	2.48	9	0.87
TR03201	42.84	44.35	3.77	1708	2.65	9	2.31	11	0.77
MR00027	67.19	68.56	9.45	4722	2.65	10	2.48	8	0.89

It is evident that for the sample of 113 segments there is little difference between the ranking performance of the *Accident Rate* and *Bayesian estimate* methods.

The last column w_1 represents the weight given to the observed accident rate (R) in the calculation of the Bayesian estimate. The w_1 values are relatively high, which could explain the similarity in the ranking performance of the *Accident Rate* and *Bayesian estimate* methods.

5.5.3 POTENTIAL ACCIDENT REDUCTION (PAR) INDEX METHOD

5.5.3.1 CONVENTIONAL APPROACH

According to McGuigan (1982) the PAR index provides a significantly better measure of the potential accident reduction, and hence the potential economic returns, than the *accident number* or the *accident rate*.

For *accident numbers* the PAR is determined as follows :

$$PAR_i = X_i - X_a (i) \quad \dots[5.17]$$

For *accident rates* the PAR index is determined as follows :

$$PAR_i = E_i(R_i - R_a) \quad \dots[5.18]$$

According to McGuigan (1982) the *PAR index* favours locations where medium risk and medium exposure prevail instead of locations where high risk and low exposure conditions prevail. According to Sayed and Rodriguez (1999) sites with a higher level of exposure are more cost effective to treat.

In Example 5.13 the total *PAR* for the ‘accident number’ sites exceeds that of the ‘accident rate’ sites. This seems to indicate that treating the ‘*accident number*’ sites could provide greater benefit than the ‘*accident rate*’ sites, assuming the cost of treating an accident rate site is equal to that of treating an accident number site.

EXAMPLE 5.15

In Table 5.19 the top ten sites, from Example 5.13, Table 5.17 is ranked according the PAR index. It is evident that the PAR method of ranking locations produces vastly different results from using the *accident rate* method.

Table 5.19 : Top 10 sites according to PAR index.

Road	Start Km	End Km	AADT	R	R Rank	PAR	PAR rank
MR00165	3.63	7.47	5530	2.45	12	13.26	1
TR00202	37.09	42.36	7300	1.48	26	10.35	2
NR00205	40.64	49.34	3063	1.49	25	7.30	3
MR00165	0	3.63	1502	4.02	3	6.53	4
NR00205	52.62	58.6	3447	1.56	22	6.18	5
NR00107	0	34.86	2526	0.91	35	5.45	6
TR03201	0	5.71	1729	2.22	15	5.33	7
NR00205	49.34	51.8	4107	2.17	17	5.27	8
TR00204	50.54	55.03	2663	1.95	19	5.27	9
NR00205	51.88	52.62	2704	7.87	1	5.21	10

5.5.3.2 EMPIRICAL BAYESIAN APPROACH

Persaud et al. (1999b) proposed that sites be ranked according to their potential safety improvement (S), where :

$$S = (m-P) \quad \dots[5.19]$$

Where

m - *The Bayesian estimate of safety*

P - *The level of safety for similar sites*

An important issue to consider is what value to use for P . Persaud et al. (1999) identified and investigated three different methods.

i) **The AADT model method**

$$S_T = (m-P_T) \quad \dots[5.20]$$

The Bayesian estimate m is calculated using the best possible regression model which incorporates all the available variables which may contribute to poor safety. P_T is based on a model that includes traffic but no treatable variables. Therefore untreatable variables such as road classification (freeway/two-lane, urban/rural) for example may be included in the model to estimate P_T . (Persaud et al. ; 1999)

The value of S in this instance is an estimate of the number of treatable accidents that are caused by all those treatable variables included in the full model.

This method will be illustrated in Example 5.16.

EXAMPLE 5.16

Persaud et al. (1999) used the following model to estimate m , (hereafter referred to as the *full* model.) :

$$P = 0.000637LQ^{0.8993}e^{-0.082157W} \quad (k = 5.82)$$

P - Injury accidents per year

L - Length of segment (km)

Q - AADT – Annual Average Daily Traffic

TW – Total pavement width

The AADT model to estimate P_T is as follows :

$$P_T = 0.000532LQ^{0.8036}$$

Assume a road segment has the following parameters :

$$L = 2 \text{ km}$$

$$Q = 4000 \text{ veh/day}$$

$$TW = 7\text{m}$$

$$X = 4 \text{ accidents per year}$$

From the *full* model $P = 2.2$ injury accidents / year and from the AADT model $P_T = 0.8$ injury accidents / year.

$$E(m) = 2.2 \text{ and } \text{VAR}(m) = (2.2)^2/5.82 = 0.83$$

$$a = 2.2/(2.2+0.83) = 0.73$$

$$m = 0.73(2.2) + (1-0.73)4 = 2.69 \text{ injury accidents/year}$$

$$S_T = 2.69 - 0.80 = 1.89 \text{ injury accidents/year}$$

ii) The BASE model method

$$S_B = (m - P_B) \quad \dots[5.21]$$

P_B is for a base condition, reflecting that what is normal can be found in the predominant values of treatable variables. According to Persaud et al. (1999) the idea is that roads are built to some desirable standard from a safety point of view. P_B is therefore estimated from a *base* model that was calibrated using those sites which represent the ideal design.

EXAMPLE 5.17

In order to estimate P_B , Persaud et al. (1999) calibrated the following model using only those segments with a lane width of 3.5m and a shoulder width of 1.8m. These magnitudes were decided on by examining the frequency of various lane width and shoulder width combinations in their total sample of 2-lane rural road segments. A review of road safety literature also revealed that these parameters constituted what could be considered safe design.

$$P_B = 0.000826LQ^{0.7448}$$

From the previous example $m = 2.69$ injury accidents / year and $Q = 4000$ veh/day.

$$P_B = 0.000826(2)(4000)^{0.7448} = 0.80 \text{ injury accidents / year.}$$

$$S = 2.69 - 0.80 = 1.89 \text{ injury accidents/ year.}$$

It can therefore be expected that increasing the total surfaced width from 7m to 10.3m will decrease the injury rate by 1.89 accidents per year.

iii) **The FULL model method**

$$S_F = m - P_F \quad \dots[5.22]$$

P is estimated from the *full* model used for the estimation of m .

S_F in this instance provides an estimate of the potential accident reduction related to those factors that were omitted from the model.

EXAMPLE 5.18

From Example 5.16 $m = 2.69$ injury accidents /year and $P = 2.2$ injury accidents / year.

$$S = 2.69 - 2.20 = 0.49 \text{ injury accidents/year.}$$

Over the 2 km segment there are factors present, treatable or untreatable, that are causing the injury accident rate to be 0.49 injury accidents / year higher than it should be. Hypothetically speaking it could be that the quality of road signs and road markings are inferior when compared to other identical segments, or the segment carries more night-time traffic, or more heavy vehicles, etc.

5.5.4 DEGREE OF DEVIATION – N_i

$$N_i = \frac{R_i}{R_{ai}} \text{ or } N_i = \frac{X_i}{X_{ai}} \quad \dots[5.23]$$

This ranking criterion considers the deviation from the expected values regardless of the accident number/rate and the level of exposure. According to Sayed and Rodriguez (1999) this index should be used in conjunction with the PAR index. They propose that the two ranking criteria be given equal weights. However if different weights are used a higher weight should be given to the PAR criterion to achieve higher cost effectiveness.

5.5.5 SEVERITY METHOD

5.5.5.1 CONVENTIONAL APPROACH

Sites identified as hazardous locations using the *EAN method* are often ranked according to their EAN values.

A *Severity Index* can be used to rank hazardous locations to favour those locations where the accidents are relatively severe.

For each selected hazardous location a *Severity Index* can be determined as follows :

$$SI = \frac{EAN}{F + SR + SL + D} \quad \dots[5.24]$$

where

F, SR, SL, D - The number of fatal, serious injury, slight injury and damage only accidents respectively.

Using the *Severity Index* as a ranking criterion will favour those sites where accidents are relatively more severe, irrespective of the number of accidents or level of exposure at the sites.

5.5.5.2 EMPIRICAL BAYESIAN APPROACH

Persaud et al.(1999a) presents a variation of the *Bayesian PAR method*, in which PAR indices are determined for each severity class and then combined with *weighting factors* to obtain a weighted PAR index.

The *weighted* Bayesian PAR method requires Safety Performance Functions for each severity class.

Let

$$PAR_{fat} = m_{fat} - P_{fat} \quad \dots[5.25]$$

$$PAR_{inj} = m_{inj} - P_{inj} \quad \dots[5.26]$$

$$PAR_{dam} = m_{dam} - P_{dam} \quad \dots[5.27]$$

where

PAR_{fat} - The potential reduction in fatal accidents.

PAR_{inj} - The potential reduction in injury accidents.

PAR_{dam} - The potential reduction in damage only accidents.

m_{fat} - The Bayesian safety estimate for fatal accidents.

m_{inj} - The Bayesian safety estimate for injury accidents.

m_{dam} - The Bayesian safety estimate for damage only accidents.

P_{fat} - The fatal accident safety estimate from a suitable SPF.

P_{inj} - The injury accident safety estimate from a suitable SPF.

P_{dam} - The damage only safety estimate from a suitable SPF.

The *weighted PAR Index* that can be used to rank and identify locations potentially suitable for remedial action is determined as follows :

$$PAR_{index} = w_{fat} * PAR_{fat} + w_{inj} * PAR_{inj} + PAR_{dam} \quad \dots[5.28]$$

EXAMPLE 5.19

A segment of 2-lane rural road, 5km long, experienced in a year 3 fatal, 5 injury and 10 damage only accidents.

The severity based Safety Performance Functions applicable to this road segment is as follows :

$$P_{fat} = L(0.00004)AADT^{1.02} \quad (k = 3.2)$$

$$P_{inj} = L(0.00008)AADT^{1.03} \quad (k = 4.0)$$

$$P_{dam} = L(0.00025)AADT^{0.98} \quad (k = 3.7)$$

Assuming that AADT = 4500 veh/day.

$$P_{fat} = 1.06 \text{ and } VAR(P_{fat}) = (1.06)^2/3.2 = 0.35 \text{ and } a = 1.06/(1.06+0.35) = 0.75$$

$$P_{inj} = 2.32 \text{ and } VAR(P_{inj}) = (2.32)^2/4 = 1.35 \text{ and } a = 2.32/(2.32+1.35) = 0.63$$

$$P_{dam} = 4.75 \text{ and } VAR(P_{dam}) = (4.75)^2/3.7 = 6.10 \text{ and } a = 4.75/(4.75+6.10) = 0.44$$

$$m_{fat} = (0.75)(1.06) + (1-0.75)(3) = 1.55$$

$$m_{inj} = (0.63)(2.32) + (1-0.63)(5) = 3.31$$

$$m_{dam} = (0.44)(4.75) + (1-0.44)(10) = 7.69$$

$$PAR_{fat} = 1.55 - 1.06 = 0.49$$

$$PAR_{inj} = 3.31 - 2.32 = 0.99$$

$$PAR_{dam} = 7.69 - 4.75 = 2.94$$

Assuming $w_{fat} = 55$ and $w_{inj} = 16$ then :

$$PAR_{index} = 55(0.49) + 16(3.31) + 7.69 = 87.6$$

5.6 PERFORMANCE OF IDENTIFICATION AND RANKING METHODS

Persaud et al. (1999b) applied the following methodology to evaluate the performance of the following ranking methods :

- Accident number (X)
- Accident Rate (R)
- Empirical Bayes (m)
- AADT model estimate ($S_T = m - P_T$)
- Base model estimate ($S_B = m - P_B$)
- Full model estimate ($S_F = m - P_F$)

Two-lane paved rural highways in Ontario, Canada were divided into 500 m non-overlapping segments. This process yielded a total of 28000 segments. The ranking of these locations were based on 3 years of data (1988 – 1990). Data for the period 1991 – 1993 were used to assess the relative performance of the ranking methods. Intersection accidents were excluded and only casualty accidents were considered.

Each of the different methods as listed above were used to identify the worst 1000 segments. This process thus yielded a total of 6 groups consisting of a 1000 sites each. Some sites might be common to all the groups and some but be unique to only one group.

In order to compare the different methods it was assumed that the group with the most *target accidents* in the subsequent period (1991 – 1993) would be the best method.

Two types of *target accidents* were considered :

- All injury accidents

- Accidents that are treatable by highway engineering methods, such as for example lane widening. An estimate of these accidents equals those that are in excess of what is normally expected on intersections of a similar class and traffic intensity.

Table 5.20 : Data for comparison of identification and ranking methods.

Method	Injury accidents on 1000 worst 0.5 km segments		
	Total accidents	Total accidents	Treatable accidents
	1988 - 90	1991-93	$\Sigma(91-93 \text{ count} - \text{AADT model estimate})$
Number	3323	1232	664
Rate	1535	336	274
Empirical Bayes	2444	1438	597
AADT model	3179	1338	681
Base model	3126	1393	673
Full Model	3278	1315	679

Source : Persaud et al. (1999b)

Using the Empirical Bayes estimate (m) as a ranking criterion appears to be the most effective method to identify and rank sites that are most likely to have accidents in the subsequent period (1991 – 1993). This is followed by the Base model method. When only treatable accidents are considered the ‘model’ methods are superior to the method based on the Empirical Bayes estimate.

Persaud et al. (1999b) conclude that if both accident measures are considered together the *Base model* method appears to be the best of the methods.

5.7 SUMMARY and CONCLUSION

This Chapter presented a number of *Conventional* and *Bayesian* methodologies to identify and rank hazardous locations or locations that when treated, would have the potential to produce good benefits in terms of accident reduction and economic returns.

Generally those *Conventional* and *Bayesian* methods that are based on a measure of *potential accident reduction* (PAR), according to McGuigan (1982) and Persaud et al. (1999a and 1999b) appear to produce satisfactory results. It has been shown how information about the *severity* of accidents at a location can be combined with the *Bayesian PAR method* to favour those sites with a good potential for *accident and severity* reduction.

CHAPTER 6

THE EVALUATION OF ROAD SAFETY REMEDIAL MEASURES

6.1 INTRODUCTION

According to Persaud (1986) the effective management of safety on a system requires sound knowledge of how the system reacts to the implementation of measures that affect safety – whether safety increases or decreases and by how much.

Hauer (1997) states that one of the main reasons for conducting evaluation studies to determine the safety effect of a treatment/s on an entity or a group of entities is to have factual guidance for the future. Hauer (1997) also states that the results of each such evaluation adds to the edifice of professional knowledge and the quality of professional advice.

The systematic and correct evaluation of road safety remedial measures is therefore a very important last step in the implementation of an accident remedial programme, since the results so obtained will add to the quality of planning and implementation of any future road safety remedial programmes.

Another important reason to have reliable information on the effect of road safety treatments is to be able to conduct economic feasibility studies. In other words, to determine how efficiently funds and resources were utilised. The overall success of a road safety remedial programme is often not measured in the reduction of accidents but in a comparison of cost and benefits. The cost component is determined by the amount of resources applied to bring about improvements in safety, while the benefit component is determined by expressing the safety effect of a treatment in financial terms using the unit accident cost in Table 1.2.

The objective of this Chapter is to present accident data analysis methodologies to determine reliable estimates of the effect that road safety remedial measures may have on the safety of an entity or a group of entities.

As with measurement, identification and ranking methodologies, evaluation methods can be divided into two categories a) Conventional methods and b) Bayesian methods.

Although the Conventional methods have been found to be less reliable than the Bayesian methods especially with regards the regression-to-mean effect (Hauer ; 1997) , for comparison purposes both the Conventional and Bayesian methods will be presented in this Chapter.

The methodologies presented in this Chapter is based on the work done by Dr Ezra Hauer in his book – “*Observational Before-and-After Studies in Road Safety*” (Pergamon – 1997).

To evaluate the performance of the *Conventional* and *Bayesian* methods and to compare these methods with one another, use will be made of the experimental data in Appendix A1. The analyses performed on this data, the results of which are used in the Chapter, are contained in Appendix A4.

6.2 CONCEPTUAL FRAMEWORK

Estimating the effect of highway and traffic engineering measures on road safety basically involves three tasks (Hauer ; 1997):

- a) *Predicting* what would have been the safety in the ‘after’ period had treatment not been applied.
- b) *Estimating* what the safety in the ‘after’ period was after treatment has been applied.
- c) Comparing the *prediction* with the *estimation* in order to estimate the improvement/decline in safety between the ‘before’ and ‘after’ periods.

6.2.1 PREDICTING ACCIDENTS

There are a number of ways to predict the expected level of safety in the 'after' period had treatment not been applied. According to Hauer (1997) the best method is the method that provides the best prediction. The quality of prediction is determined by (Hauer ; 1997) :

- a) How the method accounts for those causal factors that affect safety that are measured and the influence of which is known or can be known, such as for example *traffic flows*.
- b) How the method accounts for the remaining factors that affect safety, those that are not measured or of which the influence of safety is unknown.
- c) How the method accounts for the *regression-to-mean* effect.
- d) How the method accounts for changes in the extent of accident reporting.

6.2.2 STATISTICAL FRAMEWORK

According to Hauer (1997) let

- π - The estimated number of *target* accidents of a specific entity in the 'after' period had the entity not been treated. I.e. π is to be *predicted*.
- λ - The expected number of *target* accidents of the entity in the 'after' period. I.e. λ is to be *estimated*.

The safety effect of a treatment is determined by comparing π and λ as follows :

- a) Unbiased estimates of the reduction in the number of accidents (δ), and its variance, $VAR(\delta)$, is given by :

$$\delta = \pi - \lambda \quad \dots[6.1]$$

$$VAR(\delta) = VAR(\pi) + VAR(\lambda) \quad \dots[6.2]$$

- b) Unbiased estimates of the *index of effectiveness* (θ) and its variance, $VAR(\theta)$, are given by :

$$\theta = (\lambda/\pi)[1+VAR(\pi)/\pi^2]^{-1} \quad \dots[6.3]$$

$$VAR(\theta) = \theta^2[VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2][1 + VAR(\pi)/\pi^2]^2 \quad \dots[6.4]$$

For derivations of these expressions the reader is referred to Hauer (1997).

For a treatment to have been effective $\delta > 0$ and $\theta < 1$. The accident reduction factor :

$$ARF = 100(1-\theta). \quad \dots[6.5]$$

Equations 6.1, 6.2, 6.3 and 6.4 form the basis for all the evaluation procedures that will be presented in this Chapter. The difference between the different evaluation procedures lies in the methods used to obtain π , λ , $VAR(\pi)$ and $VAR(\lambda)$.

6.2.3 METHODOLOGICAL FRAMEWORK

The framework for the statistical analysis of any of the evaluation procedures in this chapter consists of a simple 4-step process as indicated in Table 6.1.

Table 6.1 : The 4-step process

Step 1	For $j = 1, \dots, n$ estimate $\lambda(j)$ and $\pi(j)$. $\lambda = \Sigma\lambda(j)$ $\pi = \Sigma\pi(j)$
Step 2	For $j = 1, \dots, n$ estimate $\text{VAR}\{\pi(j)\}$ and $\text{VAR}\{\lambda(j)\}$. $\text{VAR}(\lambda) = \Sigma\text{VAR}\{\lambda(j)\}$ $\text{VAR}(\pi) = \Sigma\text{VAR}\{\pi(j)\}$
Step 3	Determine δ and θ : $\delta = \pi - \lambda$ $\theta = (\lambda/\pi)[1 + \text{VAR}(\pi)/\pi^2]^{-1}$
Step 4	Determine $\text{VAR}(\delta)$ and $\text{VAR}(\theta)$: $\text{VAR}(\delta) = \text{VAR}(\pi) + \text{VAR}(\lambda)$ $\text{VAR}(\theta) = \theta^2[\text{VAR}(\lambda)/\lambda^2 + \text{VAR}(\pi)/\pi^2]/[1 + \text{VAR}(\pi)/\pi^2]^2$

Step 2 is based on the assumption that all the $\lambda(j)$'s and all the $\pi(j)$'s are mutually independent. Steps 3 and 4 are common to all the evaluation procedures that will be presented in this Chapter.

6.3 SIMPLE BEFORE AND AFTER METHODOLOGY

The *simple before-and-after* methodology method consists of comparing the observed accident count of the 'before' period (X_b) to the observed accident in the 'after' period (X_a). The observed number of 'before' accidents is therefore used as an estimate of what would have been the accident number in the after period had safety treatments not been undertaken.

The difference between two observed values X_a and X_b in an evaluation study could consist of 4 components :

a) The treatment effect.

The estimation of this effect is the primary objective of a road safety evaluation study. The *treatment effect* is that change in the level of

safety that was caused by the influence of a safety measure/s, and that alone.

b) The exposure effect

Between the 'before' and 'after' periods there may be an increase (or a decrease) in the traffic volumes. Since higher traffic volumes are associated with increased accident numbers, an increase in traffic volumes after the implementation of road safety measures could lead to a relative increase in accidents. Not accounting for the *exposure effect*, when there has been an increase in traffic volumes, could result in the underestimation of the treatment effect.

c) The trend effect

The *trend effect* is a function of a number of possible causal factors which are possibly difficult to identify and measure. Between the 'before' and 'after' periods there may have been changes in the traffic composition (i.e. more minibus taxis), changes in the driver composition (e.g. more older drivers), changes in law enforcement activity (e.g. the Arrive Alive campaign), the pedestrian numbers may have increased, the level of accident reporting may have improved etc.

d) Random effect

Accident counts are random variables which have a Poisson distribution around a long term mean. (Abbess et al. ; 1981). This long term mean will hence be referred to as the 'true' level of safety. Because of the discrete and random nature of accident data is it unlikely that an observed value will be equal to its 'true' mean. The difference between the mean and the observed value can be referred to as the *random error*. Accident remedial measures are normally applied to accident sites with a high accident number in the 'before' period. This introduces what is called *selection bias* or the *regression-*

to-mean (RTM) effect. (Abbess et al. ; 1981). It can be expected for the high observed count to decrease (towards the mean) even in the absence of any remedial measure, thereby creating a false sense of 'success'.

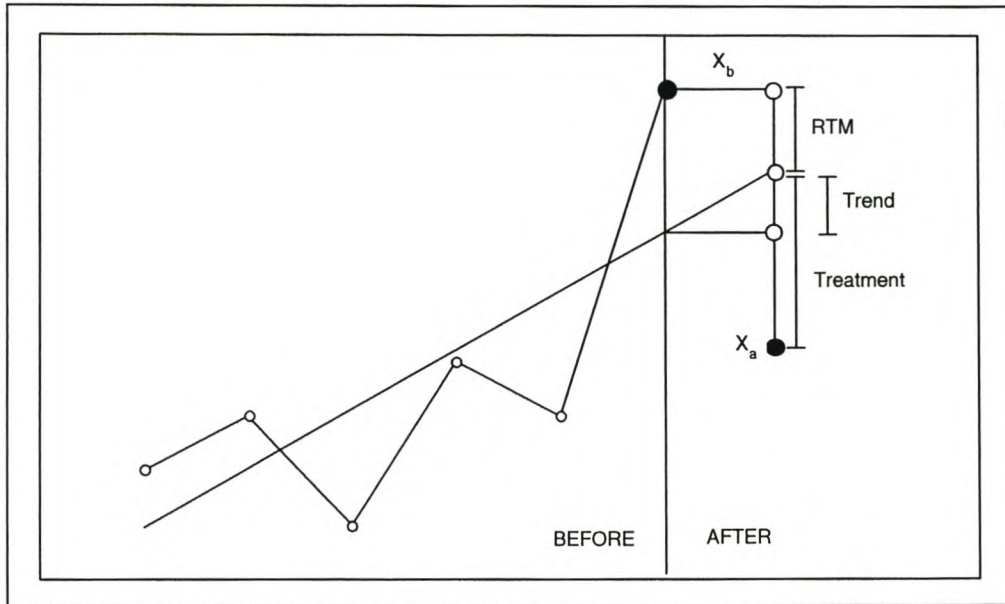


Figure 6.1 : Illustration of the RTM, treatment and trend effects

Figure 6.2 illustrates the regression-to-mean, trend and treatments effects, assuming there were no changes in the level of exposure between the 'before' and 'after' periods. If the trend in the 'before' data was not considered in predicting the expected accident number in the 'after' period the RTM effect would have been overestimated by the 'trend effect' and the treatment effect would have been underestimated by the 'trend effect'.

The following example, based on the experimental data of Appendix A1, will serve to illustrate the regression-to-mean effect.

EXAMPLE 6.1

The data in columns 'Year 2', 'Year 3', 'Year 4' and 'Year 5' of Table 6.2 are the annual average number of accidents that occurred at all those sites that during Year 1 experienced 0, 2, 4, 6 and 8 accidents respectively.

Table 6.2 : Annual average accident frequency data

Annual average frequency				
Year 1	Year 2	Year 3	Year 4	Year 5
0.00	3.67	4.33	3.50	4.44
2.00	3.83	4.06	4.07	4.06
4.00	4.07	3.80	3.89	4.35
6.00	3.91	4.05	4.13	4.18
8.00	4.96	3.89	4.36	3.82

The data in Table 6.2 is illustrated in Figure 6.2.

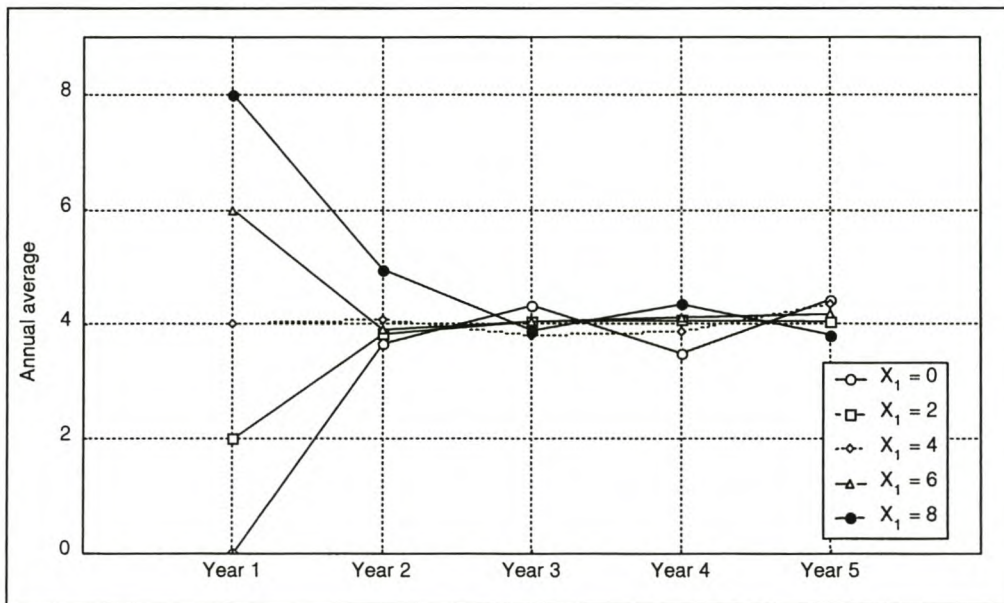


Figure 6.2 : Illustration of regression-to-mean effect

It is evident that the average annual accident frequency of all those sites that in Year 1 experienced accident frequencies above the mean (=4) 'regressed' in subsequent years towards the mean. And vice versa, annual accident frequencies of all those locations that were below the mean in Year 1 'regressed' towards the mean in subsequent years.

6.3.1 STATISTICAL ANALYSIS

Assume some treatment has been applied to a number of entities numbered 1,2,3j,.....n.

During the 'before' period the accident counts were $X_b(1), X_b(2), \dots, X_b(n)$, and the 'after' accidents were $X_a(1), X_a(2), \dots, X_a(n)$.

Since it is possible for the 'before' and the 'after' periods to differ in length from entity to entity it is necessary to define the 'ratio-of-durations' :

$$r_d(j) = T_a(j) / T_b(j) \quad \dots[6.6]$$

Where

$T_a(j)$ - Duration of after period for entity j.

$T_b(j)$ - Duration of before period for entity j.

Table 6.3 : The 4-step process for the simple before-and-after procedure

STEP 1	STEP 2
$\lambda = \sum X_a(j)$	$\text{VAR}(\lambda) = \sum X_a(j)$
$\pi = \sum r_d(j) X_b(j)$	$\text{VAR}(\pi) = \sum r_d(j)^2 X_b(j)$
STEP 3	STEP 4
$\delta = \pi - \lambda$	$\text{VAR}(\delta) = \text{VAR}(\pi) + \text{VAR}(\lambda)$
$\theta = (\lambda/\pi)[1 + \text{VAR}(\pi)/\pi^2]^{-1}$	$\text{VAR}(\theta) = \theta^2[\text{VAR}(\lambda)/\lambda^2 + \text{VAR}(\pi)/\pi^2][1 + \text{VAR}(\pi)/\pi^2]^2$

The application of this 4 step procedure is illustrated in Example 6.2.

EXAMPLE 6.2

During the first 6 months of 1998 a total of 1103 casualty accidents occurred on rural roads in the province of the Western Cape. The Arrive Alive campaign started on 1 October 1997. The following example will show the application of the simple before and after methodology in determining if there was a reduction in the casualty accident number over this period.

The following data is available for the period January to June each year:

Table 6.4 : Accident Data

Year	1994	1995	1996	1997	TOTAL
Accidents	1116	1174	1119	934	4343

$X_b = 4343; X_a = 1103 ; T_b = 4 ; T_a = 1 ; R_d = 0.25$

Table 6.5 : The 4-step procedure ; Calculations

STEP 1	STEP 2
$\lambda = 1103$	$VAR(\lambda) = 1103$
$\pi = 4343 * 0.25 = 1086$	$VAR(\pi) = 0.25^2 * 1086 = 68$
STEP 2	STEP 4
$\delta = 1086 - 1103 = -17$	$VAR(\delta) = 1103 + 68 = 1171 = (34.2)^2$
$\theta = (1103/1086) / (1 + 68/1086^2) = 1.015$	$VAR(\theta) = (1.015)^2 [1/1103 + 68/1086^2] / [1 + 68/1086^2]^2 = 0.00099 = (0.0315)^2$

An approximate 95 % confidence interval for δ is $-17 \pm 2*(34.2) = \{-85.4 ; 51.4\}$.

Note that the shorter the 'before' period the larger r_d will become. This will result in larger values for $VAR(\pi)$ and also for $VAR(\delta)$ and $VAR(\theta)$. Larger values of $VAR(\delta)$ and $VAR(\theta)$ will result in less accurate estimates of θ and δ . It will make it more difficult to detect a change in the accident number with any degree of significance.

In this particular example the calculated effect includes an exposure effect and a trend effect. It is realistic to assume that rural traffic has grown since 1994 and that there have been many other causal factors that could have influenced traffic and accident patterns over this period.

EXAMPLE 6.3

Assume a treatment has been applied to six entities e.g. intersections.

Table 6.6 : Before-and-After accident data and calculations

Entity number	Years before	Years After	X_b	X_a	$r_d(j)$	$r_d(j)X_b$	$r_d(j)^2X_b$
1	3	1	15	4	0.33	5.0	1.7
2	4	2	23	10	0.50	11.5	5.8
3	3	3	10	8	1.00	10.0	10.0
4	2	1	9	3	0.50	4.5	2.3
5	2	1	12	4	0.50	6.0	3.0
6	3	2	17	8	0.67	11.3	7.6
TOTAL				37		48.3	30.2

Table 6.7 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 37$	$VAR(\lambda) = 37$
$\pi = 48.3$	$VAR(\pi) = 30.2$
STEP 3	STEP 4
$\delta = 48.3 - 37 = 11.3$	$VAR(\delta) = 37 + 30.2 = 67.2$
$\theta = (37/48.3)/(1+30.2/48.3^2) = 0.75$	$\sigma(\delta) = 8.2$
ARF = $100(1-0.75) = 25\%$	$VAR(\theta) = (0.75)^2 / [1/37 + 30.2/48.3^2] / [1+30.2/48.3^2]^2 = 0.023$
	$\sigma(\theta) = 0.15$

The approximate 95 % confidence interval of θ is given by $0.75 \pm 2(0.15) = \{0.45 ; 1.05\}$. Since $\theta=1$ is contained in this interval there is no certainty that the treatment was indeed effective.

6.3.2 STUDY DESIGN

According to Hauer (1997) there are two important issues to consider when designing a simple *before-and-after* study :

- a) How many accidents should occur in the before period to estimate a change in safety with satisfactory precision. For example, how many

accidents should occur in the 'before' period in order to detect a 10 % reduction in the expected number of target accidents ?

b) What should be the duration of the 'before' and 'after' periods ?

Hauer (1997) derived the following expression that can be used to guide deliberations about the number of accidents required :

$$\sum X_b(j) = (\theta/r_d + \theta^2) / \text{VAR}(\theta) \quad \dots[6.7]$$

Where $\sum X_b(j)$ is the sum of the number of all expected 'before' accidents across all treated entities.

According to Hauer (1997) the following rule of thumb applies :

“The standard deviation of the estimate has to be 2 – 3 times smaller than the effect which one expects to detect.”

A factor of 2 corresponds with a 95 % degree confidence while a factor of 3 corresponds with a 99 % degree of confidence.

EXAMPLE 6.4

Consider Example 6.2. The question is how many accidents should we have in the 'before' period, to be able to detect a 10 % reduction as significant at the 95 % degree of confidence ?

$$R_d = 0.25 ; \theta = 0.9 ; \sigma(\theta) = 0.10/2 = 0.05 ; \text{VAR}(\theta) = 0.05^2$$

Substituting into Equation 6.7 gives :

$$X_b = (0.9/0.25 + 0.9^2)/0.0025 = 1764$$

In the before period we have observed 4343 accidents which implies that one would be able to detect a 10% reduction between the before and after periods. From Table 6.8 it is evident that one would not be able to detect a 5% reduction with the available information.

In the example above the required number of 'before' accidents can be reduced by increasing the duration of the 'after' period, thereby increasing r_d .

The larger the numbers in the 'before' and 'after' period the larger the statistical precision. The accident numbers can be increased in one of two ways :

- a) By increasing the number of entities for which accidents are counted, and
- b) By increasing the duration of the 'before' and 'after' periods.

However careful consideration should be given to using long 'before' and 'after' periods when there has been a trend in the accident counts over these periods. For the *simple before-and-after* study to be legitimate there should be no trend in either of the 'before' and 'after' periods. The general rule is – the longer the period the better, provided there is no time trend in the levels of safety.

Table 6.8 : Minimum expected accidents in before period : 95 % degree of confidence

r_d	Expected % reduction							
	40 %	35 %	30 %	25 %	20 %	15 %	10 %	5 %
	θ							
	0.60	0.65	0.70	0.75	0.80	0.85	0.90	0.95
0.20	84	120	177	276	464	884	2124	9044
0.25	69	99	146	228	384	733	1764	7524
0.33	54	78	116	181	306	586	1415	6050
0.40	47	67	100	156	264	506	1224	5244
0.50	39	56	84	132	224	431	1044	4484
0.60	34	49	74	116	197	380	924	3977
0.67	31	45	68	108	183	354	861	3713
0.75	29	42	63	100	171	330	804	3471
0.80	28	40	61	96	164	317	774	3344
1.00	24	35	53	84	144	280	684	2964
1.25	21	31	47	74	128	249	612	2660
1.33	20	30	45	72	124	242	595	2587
1.50	19	28	43	68	117	229	564	2457
1.67	18	27	40	65	112	219	540	2354
2.00	17	24	37	60	104	204	504	2204
2.50	15	22	34	55	96	189	468	2052
3.00	14	21	32	52	91	179	444	1951
4.00	13	19	30	48	84	166	414	1824
5.00	12	18	28	46	80	159	396	1748

Table 6.9 : Minimum expected accidents in before period : 99 % degree of confidence

r_d	Expected % reduction							
	40%	35%	30%	25%	20%	15%	10%	5%
	0.6	0.65	0.7	0.75	0.8	0.85	0.9	0.95
0.20	189	270	400	622	1046	1993	4789	20390
0.25	156	223	330	514	866	1652	3977	16963
0.33	123	176	262	409	691	1322	3190	13640
0.40	105	151	224	352	595	1141	2760	11823
0.50	88	127	189	298	505	971	2354	10109
0.60	77	111	166	262	445	857	2083	8967
0.67	71	103	154	243	413	798	1942	8370
0.75	65	95	143	225	385	744	1813	7825
0.80	63	91	137	216	370	715	1745	7539
1.00	54	79	119	189	325	630	1542	6682
1.25	47	69	105	168	289	562	1380	5997
1.33	46	67	102	163	280	546	1341	5832
1.50	43	63	96	153	265	517	1272	5540
1.67	41	60	91	146	252	494	1216	5308
2.00	37	55	84	135	234	460	1136	4969
2.50	34	50	77	124	216	426	1055	4626
3.00	32	47	72	117	204	403	1001	4398
4.00	29	43	67	108	189	375	933	4112
5.00	27	41	63	103	180	358	893	3941

EXAMPLE 6.5

Consider the following annual total accident counts recorded in the magisterial area of Hermanus on the South Coast of the Western Cape.

Table 6.10 : Annual accident data : Hermanus : 1991 to 1998

Year	1991	1992	1993	1994	1995	1996	1997	1998
Acc.	372	431	411	497	593	627	676	634

If it is decided to use 4 years of 'before' data to determine what would have been the accident level in the 'after' period had the Arrive Alive campaign not been implemented then $X_b = 2393$ and $r_d = 0.25$. According to Table 6.8 this amount of data is sufficient to detect a 10 % change in the level of safety. From the analysis below it appears as if there were a 6 % increase in the number of accidents during 1998.

Continue ...

Example 6.5 (Continued)

Table 6.11 : The 4-step procedure – calculations.

STEP 1	STEP 2
$\lambda = 634$	$\text{VAR}(\lambda) = 634$
$\pi = 0.25 * 2393 = 598$	$\text{VAR}(\pi) = 0.25^2 * 2393 = 150$
STEP 3	STEP 4
$\delta = 598 - 634 = -36$	$\text{VAR}(\delta) = 634 + 150 = 784 = 28^2$
$\theta = (634/598) / (1 + 150/598^2) = 1.060.$	$\text{VAR}(\theta) = (1.06)^2 [1/634 + 150/598^2] / [1 + 150/598^2]^2$ $= 0.00224 = 0.047^2$

Judging from Figure 6.3 it is evident that because of the trend in the accidents the 4 year average does not provide a good measure of what the level of safety would have been had the Arrive Alive campaign not been implemented.

The following analysis consider only 1 year in the before period – that is 1997.

Table 6.12 : The 4-step procedure – calculations.

STEP 1	STEP 2
$\lambda = 634$	$\text{VAR}(\lambda) = 634$
$\pi = 676$	$\text{VAR}(\pi) = 676$
STEP 3	STEP 4
$\delta = 676 - 634 = 42$	$\text{VAR}(\lambda) = 634 + 676 = 1310 = 36^2$
$\theta = (634/676) / (1 + 1/676) = 0.94$	$\text{VAR}(\theta) = (0.94)^2 [1/634 + 1/676] / [1 + 1/676]^2 = 0.0027$ $= 0.052$
ARF = 6 %	

It is evident that using a 4-year before period has produced smaller standard deviations for δ and θ than using only a 1-year before period i.e. the 4-year period measured δ and more precisely, however it measured an 'incorrect' effect more precisely. In more human terms – *it is doing the wrong thing well*. It is therefore prudent in some instances to use a shorter 'before' or 'after' period even it means a reduction in the accuracy of the estimates.

It is evident from Figure 6.3 that there is still a sizeable trend effect that remains unaccounted for i.e. the effect estimate above (- 6 %) is in all likelihood an underestimation. A comparison group could be used to account for this trend effect.

Continue ...

Example 6. (Continued)

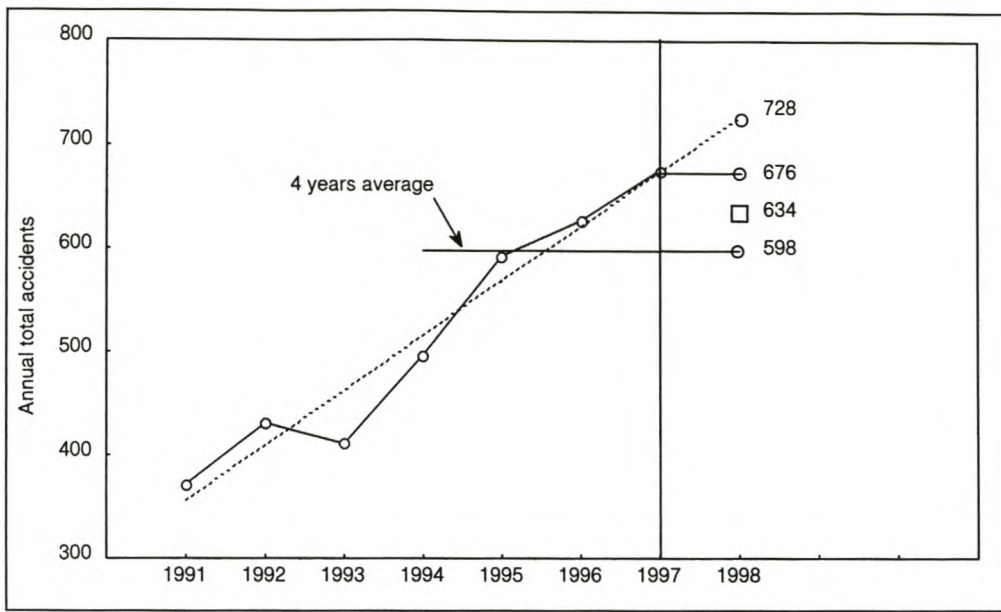


Figure 6.3 : Illustration of *trend* effect.

CASE STUDY 6.1

At the end of 1991 the municipality of Irbid, the third largest city in Jordan, implemented a number of street bumps in order to reduce traffic accident problems. Street bumps were constructed at 14 intersections of secondary streets with minor arterials, on the secondary street approaches.

A study was done by Al-Masaied (1997) to evaluate the effectiveness of the street bumps in reducing traffic accidents. He used three different methods – the simple before-and-after method, the before-and-after method with a comparison group and the Bayesian method.

The 'before' and 'after' periods are both 1 year in length.

Table 6.13 : Case study 'before' and 'after' accident data

Period	Site no.														Tot
	1	2	3	4	5	6	7	8	9	10	11	12	13	14	
Before 1991	29	16	4	8	21	34	22	13	9	19	19	23	36	16	269
After 1992	17	2	0	1	15	27	9	2	3	7	4	12	23	5	127

Table 6.14 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 127$	$VAR(\lambda) = 127$
$\pi = 269$	$VAR(\pi) = 269$
STEP 3	STEP 4
$\delta = 269 - 127 = 142$	$VAR(\delta) = 127+269 = 396 = 19.9^2$
$\theta = (127/269)/(1+1/269) = 0.47$	$VAR(\theta) = (0.47)^2[1/127+1/269]/[1+1/269]^2 = 0.0025 = 0.05^2$

An approximate 95 % confidence interval for θ is given by $0.47 \pm 2(0.05) = \{0.37 ; 0.57\}$. There appears to have been a significant reduction in the accident number between the before and after periods.

6.3.3 PERFORMANCE OF BEFORE-AND-AFTER METHODOLOGY

EXAMPLE 6.6

In order to assess the ability of this before-and-after methodology to eliminate the regression-to-mean effect and to correctly estimate the true safety effect the following analyses were applied to the hazardous locations identified by conventional methodologies (see Appendix A4) for the different collection periods. Five sets of hazardous locations were identified – each set corresponding to a different collection period.

It was assumed that the hazardous locations identified by the Conventional methodologies in Chapter 3 and Appendix A4 were treated with a treatment that would reduce the accident frequency by 20 %. The true levels of safety, which were randomly generated from a gamma distribution for each hazardous location, were reduced by 20 %. This ‘after’ true level of safety was then used as the mean to randomly generate 5 years of accident data for each site according to the Poisson distribution.

If the m_{bi} is the true level of safety at an entity, then $m_{ai} = 0.8m_{bi}$. In the before period $X_i \sim P(m_{bi})$ and in the after period $X_i \sim P(m_{ai})$.

X_b is the total number of accidents observed at all the hazardous locations during the before period, m_b the sum of all the true levels of safety at all the hazardous locations during the ‘before’ period and m_a the sum of all the true levels of safety during the ‘after’ period.

The regression-to-mean effect can be determined as follows : $RTM = 100*(1-m_b/X_b)$

Continue ...

Example 6.6 (Continued)

The expected treatment effect can be determined as follows :

$$E(\theta) = 100 \cdot (1 - m_a / X_b)$$

The simple before-and-after methodology was applied to each set of hazardous locations using different 'after' collection periods. The results of the analyses are shown in Table 6.15.

It appears from the data that as the 'before' period increases the regression-to-mean effect decreases but that the length of the 'after' period has little effect on the estimated safety effect $E(\theta)$.

Table 6.15 : Results of assessment of conventional identification and evaluation methods.

X_b	m_b	m_a	X_a	r_d	λ	π	$Var(\lambda)$	$Var(\pi)$	θ	$Var(\theta)$	σ_θ	ARF
Conventional – 1 Year												
820	503	402	414	1	414	820	414	820	0.50	0.00092	0.030	49.6
RTM = 38.7 % E(θ) = 0.51			840	2	840	1640	840	6560	0.51	0.00094	0.031	48.9
			1240	3	1240	2460	1240	22140	0.50	0.00112	0.033	49.8
			1616	4	1616	3280	1616	52480	0.49	0.00131	0.036	51.0
			2044	5	2044	4100	2044	102500	0.50	0.00160	0.040	50.4
Conventional – 2 Year												
2015	1491	1193	591	0.5	591	1008	591	252	0.59	0.00067	0.026	41.4
RTM = 26 % E(θ) = 0.408			1187	1	1187	2015	1187	2015	0.59	0.00046	0.022	41.1
			1799	1.5	1799	3023	1799	6801	0.59	0.00046	0.021	40.5
			2348	2	2348	4030	2348	16120	0.58	0.00048	0.022	41.8
			2961	2.5	2961	5038	2961	31484	0.59	0.00054	0.023	41.3
Conventional – 3 Year												
3386	2780	2224	753	0.33	753	1117	753	122	0.67	0.00065	0.025	32.6
RTM = 17.9 % E(θ) = 0.34			1496	0.66	1496	2235	1496	973	0.67	0.00039	0.020	33.1
			2255	1	2255	3386	2255	3386	0.67	0.00033	0.018	33.4
			2983	1.33	2983	4503	2983	7966	0.66	0.00032	0.018	33.8
			3769	1.66	3769	5621	3769	15489	0.67	0.00034	0.018	33.0
Conventional – 4 Year												
4589	3901	3121	790	0.25	790	1147	790	72	0.69	0.00063	0.025	31.1
RTM = 15 % E(θ) = 0.32			1575	0.5	1575	2295	1575	574	0.69	0.00035	0.019	31.4
			2380	0.75	2380	3442	2380	1936	0.69	0.00028	0.017	30.9
			3153	1	3153	4589	3153	4589	0.69	0.00025	0.016	31.3
			3951	1.25	3951	5736	3951	8963	0.69	0.00025	0.016	31.1
Conventional – 5 Year												
6352	5593	4475	1062	0.2	1062	1270	1062	51	0.84	0.00068	0.026	16.4
RTM = 11.9 % E(θ) = 0.296			1963	0.4	1963	2541	1963	407	0.77	0.00034	0.018	22.7
			2881	0.6	2881	3811	2881	1372	0.76	0.00025	0.016	24.4
			3721	0.8	3721	5082	3721	3252	0.73	0.00021	0.015	26.8
			4655	1	4655	6352	4655	6352	0.73	0.00020	0.014	26.7

6.4 ACCOUNTING FOR THE EXPOSURE EFFECT

As mentioned previously a disadvantage of the *simple before-and-after* methodology is its inability to separate the treatment effect from the effect from many other variables that changed between the ‘before’ and ‘after’ periods.

One important measurable causal factor that changes between the ‘before’ and ‘after’ period is *traffic flow* i.e. exposure. Traffic flow information is routinely collected by roads authorities for a variety of purposes. The most common measure of traffic flow for road segments is the AADT – Annual Average Daily Traffic.

It is often assumed that the exposure effect can be accounted for by using accident rates (e.g. accidents / million vehicle kilometres / year). Using accident rates in this manner is only legitimate if a direct linear relationship exists between traffic flows and accidents. As has been shown in Chapter 3 AADT based accident rates are in most cases a function of the magnitude of the AADT i.e. the Safety Performance function is non-linear.

6.4.1 STATISTICAL ANALYSIS

Let

$F(A_a)$ - The expected number of accidents when flow = A_a .

$F(A_b)$ - The expected number of accidents when flow = A_b .

The *traffic flow ratio* and its variance is given by (Hauer ; 1997) :

$$r_{ff} = F(A_a)/F(A_b) \quad \dots[6.8]$$

$$VAR(r_{ff}) = r_{ff}^2 [c_A^2 VAR(A_a) / f^2(A_a) + c_B^2 VAR(A_b) / f^2(A_b)] \quad \dots[6.9]$$

Where c_A and c_B denote the derivatives of ‘ f ’ with respect to traffic flows at A_a and A_b .

For a linear Safety Performance Function the values of r_{tf} and $VAR(r_{tf})$ is given by (Hauer ; 1997):

$$r_{tf} = A_a/A_b \quad \dots[6.10]$$

$$VAR(r_{tf}) = r_{tf}^2\{v^2(A_a) + v^2(A_b)\} \quad \dots[6.11]$$

Where

$v(A_a)$ and $v(A_b)$ are the *coefficients-of-variation* for the ‘after’ and ‘before’ counts respectively. A *coefficient-of-variation* is defined as the standard deviation of a variable divided by its mean.

For a non-linear Safety Performance Function with the following form :

$$f(A) = \alpha A^\beta \quad \dots[6.12]$$

the expressions for estimating r_{tf} and $VAR(r_{tf})$ is (Hauer ; 1997):

$$r_{tf} = (A_a/A_b)^\beta \quad \dots[6.13]$$

$$VAR(r_{tf}) = r_{tf}^2\beta^2[v^2(A_a) + v^2(A_b)] \quad \dots[6.14]$$

Table 6.16 : The 4-step procedure - accounting for the exposure effect

STEP 1	STEP 2
$\lambda = X_a$	$VAR(\lambda) = X_a$
$\pi = r_d r_{tf} X_b$	$VAR(\pi) = (r_d)^2 [(r_{tf})^2 X_b + X_b^2 VAR(r_{tf})]$
STEP 3	STEP 4
$\delta = \pi - \lambda$	$VAR(\delta) = VAR(\pi) + VAR(\lambda)$
$\theta = (\lambda/\pi)[1+VAR(\pi)/\pi^2]^{-1}$	$VAR(\theta) = \theta^2[VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2]/[1 + VAR(\pi)/\pi^2]^2$

• **THE ESTIMATION OF $v(A_a)$ and $v(A_b)$.**

The values of $v(A_a)$ and $v(A_b)$ can be obtained from Tables 2.1, 2.2 and 2.3 in Chapter 2.

EXAMPLE 6.7

Provincial Trunk Road Number 11 Section 4 connects Citrusdal in the south with Clanwilliam in the north. It forms part of the N7 route which runs from Cape Town to Namibia. During 1992 the section between kilometre values 57.01 and 76.05 was resurfaced and the road markings were repainted.

In the 4 years preceding the resurfacing i.e. 1988 to 1991 a total of 71 accidents were recorded. In the 2 years after the treatment year i.e. 1993 to 1994 a total of 15 accidents were recorded. The treatment year, 1992, was excluded from the analysis. The AADT for the 'before' period is about 1200 vehicles per day and for the after period it was about 1400 vehicles per day. The former count was obtained from a 2-day (24 hour) count and the latter from a 1-day (24 hour) count.

Table 6.17 : Data and calculations

Category : Strategic $v(A_b) = 0.177$ (From Table) $v(A_a) = 0.143$ (From Table) $T_b = 4$ $T_a = 2$ $r_d = 0.5$	$X_b = 71$ $X_a = 15$ $r_{if} = 1400/1200 = 1.167$ $VAR(r_{if}) = (1.167)^2 \{ (0.177)^2 + (0.143)^2 \} = 0.071$
---	---

Table 6.18 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 15$ $\pi = 0.5 \cdot 1.167 \cdot 71 = 41.4$	$VAR(\lambda) = 15$ $VAR(\pi) = (0.5)^2 [(1.167)^2 \cdot 71 + 71^2 \cdot 0.071] = 114$
STEP 3	STEP 4
$\delta = 41.4 - 15 = 26.4$ $\theta = (15/41.4) / (1 + 55.2/41.4^2) = 0.35$ $ARF = 100(1 - 0.35) = 65 \%$	$VAR(\delta) = 15 + 55.2 = 70.2$ $VAR(\theta) = (0.35)^2 [15/15^2 + 114/41.4^2] / [1 + 114/41.4^2]^2 =$ $\sigma(\delta) = 8.38$ $\sigma(\theta) = 0.105$

Continue ...

Example 6.7 (Continued)

If it is assumed that the Safety Performance function is non-linear and that $X \propto A^{0.8}$ then :

$$r_{tf} = (1400/1200)^{0.8} = 1.13$$

$$VAR(r_{tf}) = (1.13)^2(0.8)^2\{(0.098)^2 + (0.092)^2\} = 0.014$$

Table 6.19 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 15$	$VAR(\lambda) = 15$
$\pi = 0.5 \cdot 1.13 \cdot 71 = 40.1$	$VAR(\pi) = (0.5)^2[(1.13)^2 \cdot 71 + 71^2 \cdot 0.014] = 40.3$
STEP 3	STEP 4
$\delta = 40.1 - 15 = 25.1$	$VAR(\delta) = 15 + 40.3 = 55.3$
$\theta = (15/40.1)/(1+40.3/40.1^2) = 0.36$	$VAR(\theta) = (0.36)^2[15/15^2 + 40.3/40.1^2] / [1+40.3/40.1^2]^2 =$
$ARF = 100(1-0.36) = 64 \%$	0.011
	$\sigma(\delta) = 7.44$
	$\sigma(\theta) = 0.106$

The approximate 95 % confidence interval of the ARF is given by $64 \pm 2(10.6) = \{42.8 ; 85.8\}$. It can therefore be concluded that the resurfacing operations on TR11/4 were successful in reducing accidents.

6.5 BEFORE-AND-AFTER WITH COMPARISON GROUP METHOD

According to Hauer (1997) to increase the accuracy of prediction it is required to account for the influence of causal factors that change with time. In the previous section the changes in traffic volumes were accounted for. Whereas the change in traffic volumes can be measured, there are however numerous causal factors which are not directly measurable or which are very difficult to measure. Even if all the causal factors could be measured the analytical procedures could become very complicated and cumbersome. To collectively account for these causal factors a *comparison group* is often used.

A *comparison group*, according to Hauer (1997) is a group of entities that remained untreated and that are similar to the treated sites.

The before-and-after with comparison group procedure is based on two assumptions :

1. That the causal factors have changed from the 'before' to the 'after' period in the same manner for both the treatment and the comparison group
2. That the change in these causal factors affects the safety of the treatment and comparison group in the same way.

6.5.1 STATISTICAL FRAMEWORK

C_a The total observed number of accidents for all entities in the *comparison* group during the 'after' period.

C_b - The total observed number of accidents for all entities in the *comparison* group during the 'before' period.

μ - The total expected number of accidents for all entities in the comparison group during the 'before' period.

ν - The total expected number of accidents for all entities in the comparison group during the 'after' period.

The *comparison ratio* is defined as :

$$r_C = \nu/\mu \quad \dots[6.15]$$

The equivalent ratio for the treatment group is :

$$r_T = \pi/\chi \quad \dots[6.16]$$

If the assumptions above holds then $r_C = r_T$.

The 'odds ratio' is defined as :

$$\omega = r_C/r_T \quad \dots[6.17]$$

For each group of treatment and comparison entities there is a time series of ω 's. Any such sequence of ω 's has a mean $E(\omega)$ and a variance $VAR(\omega)$. For a comparison group to be legitimate $E(\omega) = 1$.

Table 6.20 : The 4-step procedure : Comparison group method

STEP 1	STEP 2
$\lambda = X_a$	$VAR(\lambda) = X_a$
$r_T = r_C = C_a/C_b$	$VAR(r_T) = r_T^2 [1/C_b + 1/C_a + VAR(\omega)]$
$\pi = r_T X_b$	$VAR(\pi) = \pi^2 [1/X_b + VAR(r_T)]/r_T^2$
STEP 3	STEP 4
$\delta = \pi - \lambda$	$VAR(\delta) = VAR(\pi) + VAR(\lambda)$
$\theta = (\lambda/\pi)[1+VAR(\pi)/\pi^2]^{-1}$	$VAR(\theta) = \theta^2[VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2]/[1 + VAR(\pi)/\pi^2]^2$

EXAMPLE 6.8

Assume that during 1996 a municipality converted 6 intersections from 4-way stops to signalised intersections. Assume that the comparison group consist of all non-treated 4-way stops in the municipal area. For this exercise assume $VAR(\omega) = 0$.

Table 6.21 : Treatment and comparison group data

Period	Treatment	Comparison
Before	73	307
After	59	389

Table 6.22 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 59$	$VAR(\lambda) = 59$
$r_T = 389/307 = 1.27$	$VAR(r_T) = (1.27)^2[1/307 + 1/389] = 0.0094$
$\pi = 1.27 * 73 = 92$	$VAR(\pi) = (92)^2[1/73 + 0.0094/(1.27)^2] = 165$
STEP 3	STEP 4
$\delta = 92 - 59 = 33$	$VAR(\delta) = 59 + 165 = 224 = 15^2$
$\theta = (59/92)/(1+1/92) = 0.63$	$VAR(\theta) = (0.63)^2[1/59 + 165/92^2]/[1+165/92^2]^2 = 0.014 = 0.12^2$
ARF = 37 %	S(θ) = 0.12

The approximate 95 % confidence of the ARF is given by $37 \pm 2(12) = \{13 ; 61\}$. The conversion of the 6 intersections from 4-way stops to signalised installations was therefore effective in reducing accidents.

In Example 6.8 all the entities were situated in the same municipal area.

The procedure above, as illustrated by the example, is not correct for the case where a treatment was applied to entities which do not have the same 'before' and 'after' periods and that share the same environment. Each entity therefore has to have a different comparison group. (Hauer ; 1997)

Treatments applied to different entities in different years will not have the same 'before' and 'after' periods. A year which for one entity is in the 'before' period might be in the 'after' period for another entity. It will therefore be incorrect to pool the entities because the data for the 'before' and 'after' periods will not be mutually exclusive.

The way causal factors change between the 'before' and 'after' periods and their effect of safety could be geographically sensitive. Some areas may have more rain than others, some areas may have more pedestrians, the growth in vehicle travel may be higher etc. Therefore entities that are situated in different environments cannot have a common comparison group. This would violate the basic assumptions on which the comparison group method is based on.

In such circumstances each entity should have its *own comparison ratio*. The 4-step procedure is shown in Table 6.23.

Table 6.23 : The 4 –step procedure : Entities with own comparison ratios

STEP 1	STEP 2
$\lambda(j) = X_a(j)$ $r_T(j) = r_C(j) = C_a(j) / C_b(j)$ $\pi(j) = r_T(j) X_b(j)$	$VAR\{\lambda(j)\} = X_a(j)$ $VAR\{r_T(j)\} = r_T^2(j) [1/C_b(j) + 1/C_a(j) + VAR\{\omega(j)\}]$ $VAR(\lambda) = \Sigma VAR\{\lambda(j)\}$ $VAR\{\pi(j)\} = \pi^2(j) [1/X_b(j) + VAR\{r_T(j)\}/r_T^2(j)]$ $VAR(\pi) = \Sigma VAR\{\pi(j)\}$
STEP 3	STEP 4
$\delta = \pi \cdot \lambda$ $\theta = (\lambda/\pi)[1 + VAR(\pi)/\pi^2]^{-1}$	$VAR(\delta) = VAR(\pi) + VAR(\lambda)$ $VAR(\theta) = \theta^2 [VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2] / [1 + VAR(\pi)/\pi^2]^2$

EXAMPLE 6.9

Table 6.24 : Data and calculations.

Entity	Year	X _b	X _a	C _b	C _a	VAR(ω)	r _T	VAR(r _T)	λ	π	VAR(λ)	VAR(π)
1	1983	24	18	88	95	0.007	1.08	0.034	18	26	18	47.4
2	1985	29	16	102	92	0.011	0.90	0.026	16	26	16	45.1
3	1984	30	22	71	75	0.005	1.06	0.036	22	32	22	66.5
4	1988	17	8	45	56	0.009	1.24	0.075	8	21	8	47.9
5	1992	20	11	62	70	0.007	1.13	0.048	11	23	11	44.6
6	1990	25	11	84	88	0.006	1.05	0.032	11	26	11	47.7
									86	154	86	299.3

Table 6.25 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 86$	VAR(λ) = 86
$\pi = 154$	VAR(π) = 154
STEP 3	STEP 4
$\delta = 154 - 86 = 68$	VAR(δ) = 86 + 154 = 240 = 15.5 ²
$\theta = (86/154)/(1+1/154) = 0.55$	VAR(θ) = (0.55) ² [1/86 + 1/154] / [1+1/154] ² = 0.0054 = 0.07 ²
ARF = 45 %	

In this example the calendar years covered by the 'before' and 'after' periods differed from site to site. It will be incorrect to add all the C_a and C_b values. For example consider Entities 1 and 2. For Entity 1 the year 1984 will form part of the 'after' period but for Entity 2 it will form part of the 'before' period. ΣC_a and ΣC_b are therefore not mutually exclusive.

6.5.1.1 CHOOSING A COMPARISON GROUP

It is often the case that sites are assigned to the treatment group in a non-random manner for example in terms of their poor safety records. It is therefore unlikely that the expected number of accidents in the treatment group will change in the same manner as in the comparison group.

As mentioned above a valid comparison group will have $E(\omega) = 1$. There could however be a number of possible comparison groups that could be used, possibly all with $E(\omega) = 1$. According to Hauer (1997) one should always choose the comparison group for which $1/C_b + 1/C_a + VAR(\omega)$ is the smallest.

• **How to determine VAR(ω)**

Assume information for the *treatment* as well as the *comparison* group is available for n years, and that within these n years no treatment measures have been undertaken.

Let

$$o(i+1) = [X(i)*C(i+1)] / [X(i+1)*C(i)] / [1 + 1/X(i+1) + 1/C(i)] \quad \dots[6.18]$$

$$E(o) = \sum_{i=1}^{n-1} o(i+1) \quad \dots[6.19]$$

and

$$VAR(o) = \frac{1}{n-1} \sum_{i=1}^{n-1} [E(o) - o(i+1)]^2 \quad \dots[6.20]$$

$$VAR(\omega) = VAR(o) - (1/X_a + 1/X_b + 1/C_a + 1/C_b) > 0 \text{ and } 0 \text{ otherwise.} \quad \dots[6.21]$$

EXAMPLE 6.10

Table 6.26 : Data and calculations

Year	Treatment	Comparison	O(i)
1991	60	244	
1992	63	260	0.995
1993	59	251	1.010
1994	69	274	0.916
1995	73	307	1.041
Average - E(o)			0.991
Standard deviation - $\sigma(o)$			0.053
Variance - VAR(o)			0.0028

$$VAR(\omega) = 0.0028 - (1/73 + 1/59 + 1/307 + 1/389) = -0.033$$

If $VAR(\omega) < 0$ then assume that $VAR(\omega) = 0$.

EXAMPLE 6.11

Assuming the objective of the study is to determine the reduction in the safety (if any) in the magisterial area of Cape Town between 1997 and 1998. In order for various causal factors that could have changes in magnitude and characteristics between 1997 and 1998 a comparison group has to be chosen. For this analysis three possible groups have been identified : a) The Wynberg magisterial area, b) the whole Metropolitan area which include the magisterial areas of Bellville, Kuilsriver, Somerset West, Simon’s Town, Mitchell’s Plain and Wynberg, and c) the whole Province.

The first step is to determine which of these comparison groups are the most suitable i.e. for which group does $(1/C_a + 1/C_b + VAR(\omega))$ assumes the smallest value.

Table 6.27 : Data and calculations

Year	Cape Town	Wynberg		Metro		Province	
		Acc.	o(i)	Acc.	o(i)	Acc.	o(i)
1991	15851	23528		42502		60342	
1992	17042	23999	0.949	43851	0.960	61702	0.951
1993	15090	21527	1.013	40580	1.045	58576	1.072
1994	15881	24231	1.069	45409	1.063	64846	1.052
1995	18143	24698	0.892	47823	0.922	68974	0.931
1996	18768	27679	1.083	53174	1.075	76027	1.065
1997	19603	24886	0.861	53443	0.962	73886	0.930
Average - $E(o)$		0.978		1.004		1.000	
Std. Deviation - $\sigma(o)$		0.092		0.064		0.070	
Variance - $VAR(o)$		0.009		0.004		0.005	

Table 6.28 : Data and calculations

Area	C_a	C_b	X_a	X_b	$Var(\omega)$	$1/C_b + 1/C_a + VAR(\omega)$
Wynberg	24590	24886	18827	19603	0.0088	0.0089
Metro	50167	53443	18827	19603	0.0039	0.0039
Province	75194	73886	18827	19603	0.0049	0.0049

It is evident that the smallest value of $1/C_b + 1/C_a + VAR(\omega)$ has been obtained for the Metro area. The value of $E(o)$ is also sufficiently close to 1 to make it a suitable comparison group.

Continue ...

Example 6.11 (Continued)

Table 6.29 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 18827$ $r_T = r_C = 50167/53443 = 0.94$ $\pi = 0.94 * 19603 = 18401$	$VAR(\lambda) = 18827 = (137)^2$ $VAR(r_T) = (0.94)^2 [0.0039] = 0.00345$ $VAR(\pi) = 18401^2 [1/19603 + 0.00345/(0.94)^2] = (1157)^2$
STEP 3	STEP 4
$\delta = 18401 - 18827 = -426$ $\theta = (18827/18401)/(1 + 1157^2/18401^2) = 1.019$	$Var(\delta) = 18401 + 1157^2 = 1165^2$ $VAR(\theta) = (1.019)^2 [1/18827 + 1157^2/18401^2] / [1 + 1157^2/18401^2]^2 = 0.0041$

Since $\theta > 1$ it can be concluded that there was no improvement in safety in the magisterial area of Cape Town between 1997 and 1998.

6.5.2 STUDY DESIGN

Hauer (1997) lists 5 principal choices that have to be made during the design of a *before-and-after with comparison group* procedure :

1. Select the size of the treatment group in terms of the expected number of target accidents.
2. Select the duration of the 'before' period.
3. Decide on an appropriate duration of the 'after' period.
4. Postulate what the anticipated index of effectiveness (θ) of the treatment is.
5. Select the comparison group

In order to assess the influence of the five study design considerations of the accuracy of the estimate, θ , Hauer (1997) derived the following expression for $VAR(\theta)$:

$$VAR(\theta) = (\theta/r_d + \theta^2) / X_b + \theta^2[(1/r_d + 1)C_b + VAR(\omega)] \quad \dots[6.22]$$

Tables 6.30, 6.31 and 6.32 show the minimum sample sizes required for a comparison group in order to detect a certain change in the safety level at a 95 % degree of confidence for $r_d = 0.5$, $r_d = 1$ and $r_d = 2$ respectively. It is also assumed that $var(\omega) = 0$. It is evident that the larger the anticipated *index of effectiveness* (θ) the larger the required sample size. The smaller the value of r_d i.e. the ratio of the after- to the before period the larger is the required sample size.

EXAMPLE 6.12

Consider a treatment which is expected to reduce the expected accidents by 10 %. Assuming a 1 year before and 1 year after periods, what is the smallest comparison group that one should consider in order to be able to detect a 10 % reduction ?

The expected measure of effectiveness (θ) = 0.9 and the desired standard deviation, $\sigma(\theta) = 0.05$. (assuming a 95 % degree of confidence). The minimum number of 'before' accidents is given by applying Equation 6.22.

$$X_b = (0.9+0.9^2)/(0.05)^2 = 684 \approx 700.$$

Assuming $X_b = 2000$ and assuming that the comparison group will be so good that $VAR(\omega) = 0$, Equation is applied as follows :

$$(0.05)^2 = (0.9 + 0.9^2)/2000 + (0.9)^2[2/C_b]$$

Solving gives $C_b = 984 \approx 1000$

Table 6.30 : Comparison group : Minimum sample size : $r_a = 1.0$ and $\text{var}(\omega) = 0$

Accidents Before X_b	Index of effectiveness - θ							
	0.95	0.9	0.85	0.8	0.75	0.7	0.65	0.6
	% Reduction							
	5	10	15	20	25	30	35	40
20								90
30								45
40							222	35
50							92	30
60						367	66	27
70						178	55	26
80						129	49	25
90					1080	106	45	24
100					450	92	42	21
150				3200	164	67	36	20
200				457	124	59	33	20
250				302	108	55	32	20
300			3770	246	100	53	31	19
350			1276	217	95	51	31	19
400			853	200	91	50	30	19
450			678	188	89	49	30	19
500			583	180	87	49	30	19
550			522	173	85	48	29	19
600			481	168	84	48	29	19
650			451	164	83	47	29	19
700		28350	428	161	82	47	29	19
750		7364	410	158	81	47	29	19
800		4469	395	156	80	47	29	19
850		3318	383	154	80	46	29	19
900		2700	373	152	79	46	29	18
950		2314	364	151	79	46	29	18
1000		2051	357	150	79	46	29	18
1100		1713	344	147	78	46	28	18
1200		1507	335	145	77	46	28	18
1300		1368	327	144	77	45	28	18
1400		1267	321	143	77	45	28	18
1500		1191	316	142	76	45	28	18
1600		1132	311	141	76	45	28	18
1700		1084	307	140	76	45	28	18
1800		1045	304	139	76	45	28	18
1900		1013	301	138	75	45	28	18
2000		985	299	138	75	45	28	18

Table 6.31 : Comparison group : Minimum sample size : $r_d = 0.5$ and $\text{var}(\omega) = 0$.

Accidents Before X_b	Index of effectiveness - θ							
	0.95	0.9	0.85	0.8	0.75	0.7	0.65	0.6
	% Reduction							
	5	10	15	20	25	30	35	40
20								
30								
40								1080
50								123
60							661	77
70							211	61
80							139	53
90						980	110	48
100						408	95	44
150					900	148	66	36
200					318	113	58	34
250				1846	229	98	53	32
300				758	193	91	51	31
350				533	173	86	49	30
400				436	161	83	48	30
450			8969	382	153	80	47	30
500			2779	348	147	79	47	29
550			1776	324	142	77	46	29
600			1365	306	138	76	46	29
650			1142	293	136	75	45	29
700			1001	282	133	74	45	29
750			905	274	131	74	45	28
800			835	267	129	73	45	28
850			781	261	128	72	44	28
900			739	256	127	72	44	28
950			705	251	125	72	44	28
1000			677	247	124	71	44	28
1100		19093	633	241	123	71	44	28
1200		7477	601	236	121	70	43	28
1300		4936	576	232	120	70	43	28
1400		3822	557	229	119	70	43	28
1500		3197	541	226	118	69	43	28
1600		2797	527	223	118	69	43	28
1700		2519	516	221	117	69	43	28
1800		2314	507	219	117	69	43	28
1900		2157	498	218	116	68	43	28
2000		2033	491	216	116	68	43	28

Table 6.32 : Comparison group : Minimum sample size : $r_d = 2.0$ and $\text{var}(\omega) = 0$

Accidents Before X_b	Index of effectiveness - θ							
	0.95	0.9	0.85	0.8	0.75	0.7	0.65	0.6
	% Reduction							
	5	10	15	20	25	30	35	40
20								77
30							111	30
40						490	53	23
50						129	40	20
60						86	35	19
70					378	70	32	18
80					216	61	30	17
90					162	56	28	17
100					135	52	27	16
150				313	90	43	25	15
200				200	77	40	24	15
250			1047	164	71	38	23	14
300			602	147	68	37	23	14
350			462	137	65	37	22	14
400			393	130	64	36	22	14
450			352	125	62	36	22	14
500			325	121	61	35	22	14
550		5811	306	118	61	35	22	14
600		3038	292	116	60	35	22	14
650		2164	281	114	59	35	22	14
700		1736	272	113	59	35	21	14
750		1482	265	111	59	34	21	14
800		1314	259	110	58	34	21	14
850		1194	254	109	58	34	21	14
900		1105	249	109	58	34	21	14
950		1035	245	108	58	34	21	14
1000		980	242	107	57	34	21	14
1100		897	237	106	57	34	21	14
1200		838	232	105	57	34	21	14
1300		794	229	104	57	34	21	14
1400		759	226	104	56	34	21	14
1500		732	223	103	56	34	21	14
1600		709	221	103	56	33	21	14
1700		691	219	102	56	33	21	14
1800		675	217	102	56	33	21	14
1900		661	216	102	56	33	21	14
2000		650	215	101	56	33	21	14

CASE STUDY 6.1 (continued)

For the 14 locations where street bumps were implemented in Irbid, Jordan, a total of 21 suitable comparison locations were identified. The comparison locations were selected to be similar to the locations that have been improved. Similarity included both geometric and operational characteristics. The 21 sites in the comparison group had 384 and 346 accidents in the 'before' and 'after' periods respectively.

The value of VAR(ω) has not been determined in Al-Maseied's () study and will be assumed to be equal to zero.

Table 6.33 : Treatment and comparison group accident data

$X_b = 269$	$C_b = 384$
$X_a = 127$	$C_a = 346$

Table 6.34 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 127$ $r_T = 346/384 = 0.90$ $\pi = 0.9 * 269 = 242$	$VAR(\lambda) = 127$ $VAR(r_T) = (0.90)^2 [1/384 + 1/346] = 0.0045$ $VAR(\pi) = 242^2 [1/269 + 0.0045/0.90^2] = 543 = 23^2$
STEP 3	STEP 4
$\delta = 242 - 127 = 115$ $\theta = (127/242)/(1+1/242) = 0.52$ AFR = 48 %	$VAR(\delta) = 127 + 543 = 670 = 26^2$ $VAR(\theta) = (0.52)^2 [1/127+543/242^2] [1+543/242^2]^2 = 0.0046$ $\sigma(\theta) = 0.068$

The 95 % confidence interval for ARF is approximately {34 % ; 62%}

According to Al-Maseied the mean number of accidents recorded at the 14 treatment sites in the before period is 19.2 accidents per site, while the mean number recorded at the 21 comparison group sites was 18.3 accidents per site. It is thus evident that on average the treatment sites and the comparison sites had similar levels of safety. When this is the case the regression-to-mean effect can be accounted for by the before-and-after with comparison group procedure.

The regression-to-mean effect can in this case be estimated by the difference between the result obtained from the simple study and the results obtained from the comparison group study i.e. RTM = 53 – 48 = 5%.

6.6 THE EMPIRICAL BAYES APPROACH

The conventional before-and-after procedures (with or without a comparison group) assumes that the observed accident numbers X_a and X_b are estimates of the true level of safety in the 'before' and 'after' periods i.e. m_b and m_a .

This assumption makes the conventional methods prone to the regression-to-mean effect as was illustrated in the results shown in Table 6.15.

According to Abbess et al. (1981) one of the advantages of the Empirical Bayesian approach is that the method is able to eliminate the regression-to-mean effect, provided that the prior estimates $E(m)$ are sufficiently reliable.

6.6.1 THE SIMPLE BEFORE-AND-AFTER PROCEDURE

6.6.1.1 THE COMPOSITE METHOD

The composite method is suitable to determine the average effect of an treatment when applied to a number of entities.

Some treatment has been applied to a number of entities numbered $1, 2, 3, \dots, j, \dots, n$. During the before period the accident counts were $X_b(1), X_b(2), \dots, X_b(n)$ and during the 'after' period the accident counts were $X_a(1), X_a(2), \dots, X_a(n)$. The 'before' duration and the 'after' duration may differ from entity to entity.

For each entity there is a reference population, although several treated entities may have the same reference group. For each reference population the *prior* moments $E\{m(j)\}$ and $VAR\{m(j)\}$ are known.

The *weights* for each entity is calculated as follows :

$$\alpha(j) = [1 + VAR\{m(j)\}/E\{m(j)\}]^{-1} \quad \dots[6.23]$$

$$E\{m_b | X_b(j)\} = \alpha(j)E\{m(j)\} + [1-\alpha]X_b(j) \quad \dots[6.24]$$

$$VAR\{m_b | X_b(j)\} = [1-\alpha(j)] E\{m_b | X_b(j)\} \quad \dots[6.25]$$

Table 6.35 : The 4-step procedure : Bayesian before-and-after method

STEP 1	STEP 2
$\lambda = \Sigma X_a(j)$ $\pi = \Sigma r_d(j) \pi(j)$	$VAR(\lambda) = \Sigma X_a(j)$ $VAR(\pi) = \Sigma r_d(j)^2 VAR\{\pi(j)\}$
STEP 3	STEP 4
$\delta = \pi - \lambda$ $\theta = (\lambda/\pi)[1+VAR(\pi)/\pi^2]^{-1}$	$VAR(\delta) = VAR(\pi) + VAR(\lambda)$ $VAR(\theta) = \theta^2[VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2]/[1 + VAR(\pi)/\pi^2]^2$

From Step 2 in Table 6.35 the question may arise as to why $\Sigma X_a(j)$ is used as an estimate of the true number of accidents in the ‘after’ period and not $\Sigma m_a(j) | X_a(j)$ knowing that $m_a | X_a$ is a more reliable estimate of safety than X_a ?

The answer is that if all treated sites in the ‘after’ period also forms the reference group for the ‘after’ period then it can be shown, as follows, that for the treated sites $\Sigma X_a(j) = \Sigma m_a(j) | X_a(j)$

Assume that of the n sites that formed the reference group in the ‘before’ period a total of k were treated.

$$\sum_{j=1}^k E[m(j) | X(j)] = \sum_{j=1}^k \{aE(m) + (1-a)X(j)\} \quad \dots[6.26]$$

$$\sum_{j=1}^k E[m(j) | X(j)] = a.k.E(m) + (1-a) \sum_{j=1}^k X(j) \quad \dots[6.27]$$

$$k.E(m) = \sum_{j=1}^k X(j) \quad \dots[6.28]$$

$$\sum_{j=1}^k E[m(j) | X(j)] = \sum_{j=1}^k X(j) \quad \dots[6.28]$$

Should the treated sites be too few to form a reference group on their own $E[m_a|X_a(j)]$ could be estimated for each site using the methods described in Chapter 4.

CASE STUDY 6.1 (continued)

In the evaluation of the installation of speed humps in the city of Irbid in Jordan, Al-Masaeid (1997) also used the Empirical Bayes approach to estimate the safety effect of speed hump installations. For each treated site the *method-of-moments* (see Chapter 4) was used to calculate $m_a|X_a$ and $m_b|X_b$ i.e. for both the ‘before’ and ‘after’ period. The study by Al-Masaeid (1997) did not provide enough information for the estimation of $VAR(m_a|X_a)$ and $VAR(m_b|X_b)$ and this example will be confined to Step 1 and Step 3 of the four step procedure.

Table 6.36 : Before-and-After Bayesian estimates

Location	Expected Accident Frequency	
	$m_b X_b$	$m_a X_a$
1	22.35	14.55
2	14.71	5.46
3	7.65	4.24
4	10.00	4.85
5	17.65	14.55
6	25.29	20.67
7	18.24	9.70
8	12.94	5.45
9	10.59	6.06
10	16.47	8.49
11	16.47	6.67
12	18.82	11.52
13	26.47	18.18
14	14.71	7.27
TOTAL	232.36	136.60

The (biased) measure of effectiveness $\theta = (136.60/232.36) = 0.59$ and ARF = 41 %. The simple before-and-after methodology indicated an ARF of 53 %. The regression-to-mean effect is thereforee approximately $53 - 41 = 12$ %. The before-and-after with comparison group methodology estimated an ARF of 43 % which is relatively similar the Empirical Bayes estimate. The reason for this result is that the sites comprising the comparison group shared similar levels of safety as the 14 sites in the treatment group.

6.6.1.2 PERFORMANCE OF BAYESIAN BEFORE AND AFTER METHODOLOGY

EXAMPLE 6.13

In order to assess the ability of this before-and-after methodology to eliminate the regression-to-mean effect and to correctly estimate the true safety effect the following analyses were applied to hazardous locations identified by the Bayesian methodology for the different collection periods. Five sets of hazardous locations were identified – each set corresponding to a different collection period

It was assumed that the hazardous locations identified by the Bayesian methodology in Appendix B4 were treated with a treatment that would reduce the accident frequency by 20 %. The true levels of safety, which were randomly generated from a gamma distribution for each hazardous location, were reduced by 20 %. This ‘after’ true level of safety was then used as the mean to randomly generate 5 years of accident data for each site according to the Poisson distribution.

If the m_{bi} is the true level of safety at an entity, then $m_{ai} = 0.8m_{bi}$. In the before period $X_i \sim P(m_{bi})$ and in the after period $X_i \sim P(m_{ai})$.

X_b is the total number of accidents observed at all the hazardous locations during the before period, m_b the sum of all the true levels of safety at all the hazardous locations during the ‘before’ period and m_a the sum of all the true levels of safety during the ‘after’ period.

The regression-to-mean effect can be determined as follows :

$$RTM = 100*(1-m_b/X_b)$$

The expected treatment effect can be determined as follows :

$$E(\theta) = 100*(1-m_a/X_b)$$

The simple before-and-after Bayesian methodology was applied to each set of hazardous locations using different ‘after’ collection periods. The results of the analyses are shown in Table 6.37.

Continue ...

Example 6.13 (Continued)

Table 6.37 : Assessment of simple before-and-after Bayesian methodology

X_b	M_b	m_a	r_d	λ	π	$VAR(\lambda)$	$VAR(\pi)$	θ	$\sigma(\theta)$	ARF
Bayesian – 1 Year										
68	65	52	1	58	68	58	68	0.84	0.149	15.6
RTM = 4.7 % E(θ) = 0.76			2	108	135	108	542	0.77	0.148	22.6
			3	161	203	161	1829	0.76	0.163	24.1
			4	215	271	215	4335	0.75	0.179	25.1
			5	271	339	271	8468	0.75	0.193	25.5
Bayesian – 2 Year										
314	316	253	1	262	314	262	314	0.83	0.069	16.8
RTM = -0.7 % E(θ) = 0.81			2	530	627	530	2510	0.84	0.076	16.1
			3	780	941	780	8470	0.82	0.085	17.9
			4	1021	1255	1021	20078	0.80	0.093	19.7
			5	1290	1569	1290	39214	0.81	0.103	19.1
Bayesian – 3 Year										
502	512	410	1	411	502	411	502	0.82	0.054	18.3
RTM = -2.0 % E(θ) = 0.82			2	830	1004	830	4017	0.82	0.059	17.7
			3	1242	1506	1242	13557	0.82	0.067	18.0
			4	1636	2008	1636	32134	0.81	0.074	19.2
			5	2062	2510	2062	62762	0.81	0.082	18.7
Bayesian – 4 Year										
634	642	513	1	518	634	518	634	0.82	0.048	18.4
RTM = -1.2 % E(θ) = 0.81			2	1052	1268	1052	5073	0.83	0.053	17.3
			3	1589	1902	1589	17121	0.83	0.061	16.9
			4	2099	2536	2099	40582	0.82	0.067	17.8
			5	2633	3170	2633	79262	0.82	0.074	17.6
Bayesian – 5 Year										
713	724	579	1	684	713	684	713	0.96	0.051	4.2
RTM = -1.5 % E(θ) = 0.81			2	1283	1427	1283	5707	0.90	0.054	10.3
			3	1894	2140	1894	19260	0.88	0.060	11.9
			4	2451	2853	2451	45654	0.85	0.066	14.6
			5	3038	3567	3038	89168	0.85	0.072	15.4

It is evident from the results in Table 6.37 that the Bayesian methodology was effective in eliminating the regression-to-mean effect completely. There appears to be little difference in the influence of the length of the ‘before’ and ‘after’ periods in estimating the safety effect.

6.6.2 BEFORE-AND-AFTER WITH COMPARISON GROUP METHOD

Treatments may be applied to entities that do not have the same ‘before’ and ‘after’ periods and environment. Each entity therefore has to have a separate comparison group.

Table 6.38 : The 4-step procedure : Bayesian before-and-after with comparison group

STEP 1	STEP 2
$\lambda(j) = X_a(j)$ $r_T(j) = r_C(j) = C_a(j) / C_b(j)$ $\pi(j) = r_T(j)\lambda(j)$	$VAR\{\lambda(j)\} = X_a(j)$ $VAR\{r_T(j)\} = r_T^2(j) [1/C_b(j) + 1/C_a(j) + VAR\{\omega(j)\}]$ $VAR(\lambda) = \Sigma VAR\{\lambda(j)\}$ $VAR\{\pi(j)\} = \pi^2(j) [VAR\{\lambda(j)\}/\lambda(j)^2 + VAR\{r_T(j)\}/r_T^2(j)]$ $VAR(\pi) = \Sigma VAR\{\pi(j)\}$
STEP 3	STEP 4
$\delta = \pi \cdot \lambda$ $\theta = (\lambda/\pi)[1 + VAR(\pi)/\pi^2]^{-1}$	$VAR(\delta) = VAR(\pi) + VAR(\lambda)$ $VAR(\theta) = \theta^2 [VAR(\lambda)/\lambda^2 + VAR(\pi)/\pi^2] [1 + VAR(\pi)/\pi^2]^2$

6.6.3 THE MULTIVARIATE REGRESSION ‘TIME-SERIES’ APPROACH

In Chapter 4 a methodology is presented to estimate the level of safety at an entity assuming that this level of safety changes from year to year. Hauer (1997) proposed this methodology to overcome two basic impediments of the ‘classical’ Empirical Bayes approach, namely :

- a) That over the study period the ‘true’ level of safety remains unchanged, and
- b) That the study period remains fixed and that information outside of the study period has no value.

Hauer’s (1997) methodology allows the estimation of the safety at a location i for each of Y years (i.e. $m_{i,y}$), where Y is the period for which information is available. The objective of this section is to show how Hauer’s (1997) methodology can be applied to predict what would have been the expected accident frequencies $m_{i,Y+1}$, $m_{i,Y+2}$, $m_{i,Y+Z}$ in the ‘after’ years had the treatment not been applied. Where Z = the number of years in the ‘after’ period for which information is available.

6.6.3.1 THEORETICAL FRAMEWORK

Assume we have an entity i with Y years of covariate and accident information available in the 'before' period and similar information for Z years in the 'after' period. Assume the treatment has been applied at the end of year Y . Since the objective is to determine the number of accidents that would have occurred had treatment not taken place the same reference group should be used for the 'before' and 'after' periods.

These prediction models can be used to estimate $E(m_{i,1}), \dots, E(m_{i,Y}), E(m_{i,Y+1}), \dots, E(m_{i,Y+Z})$ from which C_{ij} can be determined as follows :

$$C_{i,j} = \frac{E(m_{i,j})}{E(m_{i,1})} \quad \dots[6.29]$$

where $j = 1, 2, 3, \dots, Y+Z$.

The predicted number of accidents for each year in the 'after' period can be estimated from Eqn. 6.30 :

$$m_{i,Y+z} = C_{i,Y+z} m_{i,1} \quad \dots[6.30]$$

$$VAR(m_{i,Y+z}) = C_{i,Y+z}^2 VAR(m_{i,1}) \quad \dots[6.31]$$

where $z = 1, 2, \dots, Z$.

The value of $m_{i,1}$ can be determined from Equation 4.57 and $VAR(m_{i,1})$ from Equation 4.58.

EXAMPLE 6.14

In Example 4.9 the number of accidents on a 2km segment of road was estimated for each of 6 years. Assume now that at the end of year 6 this segment of road underwent major rehabilitation. The road which had gravel shoulders to start of with now has surfaced shoulders. For the after period 6 years worth of information is available. Because of improvements to the road the traffic volumes increased significantly over the 'after' period.

The first step in the evaluation process is to determine what would have been the total number of accidents in the 'after' had the road not been improved, given the increased traffic volumes.

In Example 4.9 there was a reference group, consisting of a sufficient number of 2-lane road segments with gravel shoulders, that was used to develop accident prediction models for each of the 6 years in the before period. A similar reference group consisting of 2-lane road segments with gravel shoulders can be used to develop accident prediction models for each of the 6 years in the after period. For all these models assume $\beta = 0.78$. The α_{y+z} values are indicated in the table below.

Table 6.39 : Data and calculations

J	α_j	$Q_{i,j}$	$E(m_{i,j})$	$C_{i,j}$	$m_{i,j}$	$VAR(m_{i,j})$
1	0.00271	1250	1.41	1.000	2.49	0.264
7	0.00283	1410	1.62	1.15	2.86	0.348
8	0.00277	1480	1.65	1.17	2.91	0.360
9	0.00269	1502	1.62	1.15	2.85	0.347
10	0.00280	1493	1.67	1.19	2.96	0.372
11	0.00291	1520	1.77	1.25	3.12	0.414
12	0.00278	1515	1.68	1.19	2.97	0.376

$$\pi_i = \sum_{z=1}^6 m_{p(i,z)} = 17.66 \text{ and } VAR(\pi_i) = \sum_{z=1}^6 VAR(m_{p(i,z)}) = 2.22$$

The second step in the evaluation process is to determine the actual number of accidents that occurred in the 'after' period. Assume the observed accident frequencies in the 'after' period are – { 2, 1, 2, 0, 3, 3}.

In order to develop prediction models for each of the 6 years in the after period a reference group consisting of 2-lane segments with surfaced shoulders, each with 6 years' worth of accident data and traffic flow information, is required. Assume accident models have been fitted to the data for each year and that the β parameter for all 6 models is 0.85.

Continue ...

Example 6.14 (Continued)

The α parameters for each year are indicated in the Table below. Assume also that the k value of the model for the first year of the 'after' period = 5.5.

From Equation 4.57 and 4.58 :

$$m_{a(i)} = (5.5 + 11)/(5.5/1.95 + 5.96) = 1.88$$

$$VAR[m_{a(i,1)}] = 1.88/(5.5/1.95 + 5.96) = 0.214$$

Table 6.40 : Data and calculations

j	α_j	$Q_{i,j}$	$E[m_{a(i,j)}]$	$C(i,j)$	$X_{i,j}$	$m_{a(i,j)}$	$VAR[(m_{a(i,j)})]$
7	0.00205	1410	1.95	1.00	2.00	1.88	0.214
8	0.00193	1480	1.91	0.98	1.00	1.84	0.206
9	0.00185	1502	1.86	0.95	2.00	1.79	0.194
10	0.00197	1493	1.97	1.01	0.00	1.90	0.218
11	0.00201	1520	2.04	1.05	3.00	1.96	0.234
12	0.00188	1515	1.90	0.97	3.00	1.83	0.203
TOTALS				5.96	11.00	11.21	1.270

$$\lambda_i = \sum_{j=1}^6 m_{a(i,j)} = 11.21 \text{ and } VAR(\lambda_i) = \sum_{j=1}^6 VAR[m_{a(i,j)}] = 1.27$$

Now applying the 4 step procedure :

Table 6.41 : The 4-Step procedure - calculations

STEP 1	STEP 2
$\lambda = 11.21$	$VAR(\lambda) = 1.27$
$\pi = 17.66$	$VAR(\pi) = 2.22$
STEP 3	STEP 4
$\delta = 17.66 - 11.21 = 6.45$	$VAR(\delta) = 1.27 + 2.22$
$\theta = (11.21/17.66)[1+2.22/17.66^2]^{-1} = 0.63$	$VAR(\theta) = (0.63)^2[1.27/11.21^2 +$
ARF = 100(1-0.63) = 37 %	$2.22/17.66^2]/[1+2.22/17.66^2]^2 = 0.00674$
	$\sigma(\theta) = 0.082$

6.6.4 ALLOWING FOR CHANGES IN EXPOSURE

Improving the safety of a location may cause traffic volumes to increase above what is normal. The simple before-and-after Bayesian methodology estimates what the number of accidents would have been in the 'after' period using traffic volumes applicable to the 'before' period. This is an acceptable approach if the policy is to consider the accidents caused by additional traffic, generated by a treatment, to be included in the estimation of the safety effect. The assumption is that had the treatment not been undertaken the traffic flows would not have increased and the additional accidents as result thereof would not have occurred.

This approach however do not provide accurate estimates of the true (gross) safety effect of a treatment. To obtain such a 'gross' estimate it is required to estimate the number of accidents that would have occurred in the before period if flows were equal to those observed in the 'after' period i.e. after treatment has been applied. This estimate then becomes what would have occurred in the after period had no treatment been undertaken.

To illustrate these concepts the following example is presented ;

EXAMPLE 6.15

Consider a section of road which during the before period had an accident rate of 2 acc/mvkm and an exposure of 2 mvkm. The road was then upgraded to the extent that the accident rate decreased to 1 acc/mvkm and the exposure increased to 3 mvkm.

The number of accidents in the before period = 4. The simple before-and-after methodology would assume that had the road not been upgraded a total of 4 accidents would have occurred during the after period. This would give a safety effect $3/4 = 0.75$. However if the section of road in the 'before' period carried the same amount of traffic as in the 'after' period a total of 6 accidents would have occurred, and since only 3 accidents occurred in the after period (at the same level of exposure) the safety effect $\theta = 3/6 = 0.5$.

The multivariate regression ‘time-series’ method can be adapted to allow for changes in traffic volumes between the ‘before’ and ‘after’ period.

Assume we have a fixed before duration of t_b years and a fixed after period of t_a years. In these periods X_b and X_a accidents were observed respectively. Also assume that during the before and after periods the traffic flows were Q_b and Q_a respectively.

Using an appropriate regression model, based on a reference group related to the entity’s ‘before’ condition, the regression estimate of the total annual number of target accidents for the ‘before’ ($E[\bar{m}_b]$) and ‘after’ ($E[\bar{m}_p]$) periods can be determined, where $E(m_b) = f(Q_b)$ and $E(m_p) = f(Q_a)$.

The expected annual number of accidents (\bar{m}_b) can be determined from the following Equations which can be derived from Equations 4.57 and 4.58 ;

$$\bar{m}_b = \frac{k + X_b}{\frac{k}{E(\bar{m}_b)} + t_b} \quad \dots[6.32]$$

$$VAR(\bar{m}_b) = \frac{\bar{m}_b}{\frac{k}{E(\bar{m}_b)} + t_b} \quad \dots[6.33]$$

Since the objective is to determine the total number of accidents that would have occurred in the ‘after’ period had the treatment not been undertaken it is necessary to determine the total expected number of accidents in a ‘before’ period equal in length to the ‘after’ period :

$$m_b = t_a \bar{m}_b \quad \dots[6.34]$$

$$VAR(m_b) = t_a^2 VAR(\bar{m}_b) \quad \dots[6.35]$$

To allow for changes in traffic flows between the 'before' and 'after' periods Hauer (1997) recommends using a correction factor (C) which is estimated as follows :

$$C = \frac{E(\bar{m}_a)}{E(\bar{m}_b)} \quad \dots[6.36]$$

The total number of accidents that would have occurred in the after period had the treatment not been undertaken (m_p) and its variance can be determined as follows :

$$m_p = Cm_b \quad \dots[6.37]$$

$$VAR(m_p) = C^2 VAR(m_b) \quad \dots[6.38]$$

For the 4 step process :

$$\pi = \sum m_p \quad \dots[6.39]$$

$$VAR(\pi) = \sum VAR(m_p) \quad \dots[6.40]$$

CASE STUDY 6.2

This case study is based on research done by Persaud et al. (2000) on crash reductions following the installation of roundabouts in the United States.

In the states of California, Colorado, Florida, Kansas, Maine, Maryland, South Carolina and Vermont a total of 24 intersections were converted to modern roundabouts between 1992 and 1997. Of the 24 intersections, 20 were previously controlled by stop signs and 4 were controlled by traffic signals. Fifteen (15) of the roundabouts were single-lane circulation designs and 9 had multi-lane designs.

For each intersection accident data were obtained for the periods before and after conversion. The construction period and the 1st month after conversion were excluded from the analysis. The lengths of the before and after periods varied in accordance with available accident data. In no case was a period shorter than 15 months.

In order to apply the multivariate regression method to estimate safety in the 'before' period a number of regression models were assembled. Bonneson and McCoy's (1993) model for rural stop controlled intersections were used to estimate the safety at those stop controlled intersections in rural areas that were converted to modern roundabouts. New models were calibrated for stop controlled urban intersections. In order to calibrate these models the reference group consisted of urban stop controlled intersections that were not converted to roundabouts.

In order to illustrate the methodology applied by Persuad et al. (2000) only one intersection will be considered. The intersection in question is situated in Anne Arundale County, Maryland, and was converted from a rural stop controlled intersection to a modern 1 lane roundabout in 1994. The information pertaining to this intersection is as follows :

Table 6.42 : Before and after data

Data Description	Before	After
Years of accident data	4.67	3.17
Accident Count	34	14
AADT on major approaches	10 654	11 956
AADT on minor approaches	4 691	5 264

The model developed by Bonneson and McCoy (1993) for rural stop controlled intersections is as follows :

$$Y = 0.000379Q_{maj}^{0.256}Q_{min}^{0.831} \quad k = 4.0$$

where

Y - The expected number of all accidents per year

Q_{maj} - AADT on major road approaches

Q_{min} - AADT on minor road approaches

The first step is to determine m_p and $VAR(m_p)$ – the expected accident frequency and its variance in the 'after' period

The expected number of accidents in the 'before' period $E(m_b) = 0.000379(10654)^{0.256}(4691)^{0.831} = 4.58 \text{ acc / year}$ and $VAR(m_b) = 4.58^2/4 = 5.25$

The values of \bar{m}_b and $VAR(m_b)$ are determined from Equations 6.32 and 6.33 as follows :

$$\bar{m}_b = (4+34)/(4/4.58 + 4.67) = 6.86 \text{ acc/year}$$

$$\text{VAR}(\bar{m}_b) = 6.86/(4/4.58 + 4.67) = 1.238$$

The next step is to multiply \bar{m}_b by the length of the after period :

$$m_b = 6.86 * 3.17 = 21.74 \text{ and } \text{VAR}(m_b) = (3.17)^2(1.238) = 12.441$$

In order to estimate m_p the differences in the AADT's between the before and after periods need to be considered.

The expected number of accidents per year in the after period had the intersection not been converted to a roundabout = $E(m_p) = 0.000379(11956)^{0.256}(5264)^{0.831} = 5.19 \text{ acc/year}$.

$$C = 5.19/4.58 = 1.133$$

$$m_p = C(m_b) = 1.133(21.74) = 24.63$$

$$\text{VAR}(m_p) = (1.133)^2(12.441) = 15.97$$

In order to determine the improvement in safety at rural stop controlled intersections the same procedure as described above was applied to all rural stop controlled intersections which were converted to roundabouts.

The results of the analysis for the 5 rural stop controlled intersections in the study are indicated in the Table below :

Table 6.43 : Results of study

X_a	M_p	$\text{VAR}(m_p)$
14	36.71	30.63
14	24.62	15.95
2	14.38	9.40
10	14.33	8.55
4	15.16	6.76
Sum = 44	Sum = 105.19	Sum = 71.29

The 4 step procedure can be applied as follows :

Table 6.44 : The 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 44$	$\text{VAR}(\lambda) = 44$
$\pi = 105.19$	$\text{VAR}(\pi) = 71.29$
STEP 3	STEP 4
$\delta = 105.19 - 44 = 61.19$	$\text{VAR}(\delta) = 44 + 71.29 = 115.29$
$\theta = (44/105.19)/[1+71.29/105.19^2]$ =0.416 .	$\text{VAR}(\theta) = (0.416)^2[1/44+71.29/105.19^2]/$ $[1+71.29/105.19^2]^2 = 0.005$
ARF = $100(1-0.416) = 58 \%$	$\sigma(\theta) = 0.07$

The approximate 95 % confidence interval of the ARF is $58 \pm 2(7) = \{34 ; 74\}$. It can therefore be concluded that the conversion of the rural stop controlled intersections in the study group to roundabouts was effective in improving safety.

6.7 THE TREATMENT EFFECT

In the *Conventional* and *Bayesian* procedures discussed thus far entities were lumped together and were treated as if they were one ‘composite entity’. Only the sum of the accident counts of all the entities was used to estimate the safety effect associated with a treatment. Consider the simple before-and-after procedure applied to the data of Case Study 6.1. It was estimated that the average effect of the speed humps was a 53 % reduction with a standard deviation of 5%. This estimated effect has been determined by considering all 12 sites together. The reduction of 53 % does not apply to any of the sites, nor does the 5% standard deviation measure how the safety effect varies from site to site. It represents the uncertainty surrounding the estimated value of the average effect.

When a treatment is applied to a number of sites the effect of this treatment will in all likelihood vary from site to site. For these different effects a mean value as well as a variance can be determined. Should the variance be small one could expect the treatment to have approximately the same effect on all

entities, and should the variance be large the effect of a treatment would be difficult to predict.

The objective of the following section is to show how the standard 4-step procedure can be expanded to estimate both the mean and the variance of the safety effect.

6.7.1 THEORETICAL FRAMEWORK

Assume the same treatment has been applied to n entities that have indices of effectiveness denoted as $\theta(1), \theta(2), \dots, \theta(n)$.

The average effect is given by :

$$\bar{\theta} = \sum \theta(j) / n \quad \dots [6.41]$$

The sample variance is given by :

$$s^2(\bar{\theta}) = \sum [\theta(j) - \bar{\theta}]^2 / (n-1) \quad \dots [6.42]$$

The expected effect $E(\theta)$ is determined as follows :

$$E(\theta) = \bar{\theta} \quad \dots [6.43]$$

The procedure to determine $VAR(\theta)$ is as follows (Hauer ; 1997):

- a) Apply the single entity 4-step procedure to estimate $\theta(j)$ and $VAR\{\theta(j)\}$ for each entity.
- b) Using the estimates $\theta(j)$ compute the sample variance.
- c) Determine the average variance for all the entities : $avg(V) = \sum VAR\{\theta(j)\} / n$
- d) $VAR(\theta)$ can be estimated by : $VAR(\theta) = s^2(\bar{\theta}) - avg(V)$

Table 6.45 : The 5-step procedure – estimating treatment effects

STEP 1	For $j = 1$ to n estimate $\lambda(j)$ and $\pi(j)$. Estimate λ and π .
STEP 2	For $j = 1$ to n estimate $\text{VAR}\{\lambda(j)\}$ and $\text{VAR}\{\pi(j)\}$ Estimate $\text{VAR}(\lambda)$ and $\text{VAR}(\pi)$
STEP 3	Estimate δ and θ
STEP 4	Estimate $\text{VAR}(\delta)$ and $\text{VAR}(\theta)$
STEP 5	Determine $E(\theta)$ and $\text{VAR}(\theta)$

CASE STUDY 6.1 (continued)

Because of zero accidents in the ‘after’ period, $\text{VAR}\{\theta(j)\}$ for Site number 3 is not defined and has this site has been left out of this analysis.

Table 6.46 : Data and calculations

Site	$\pi(j)$	$\lambda(j)$	$\delta(j)$	$\theta(j)$	ARF(j) [%]	$\text{VAR}\{\delta(j)\}$	$\text{VAR}\{\theta(j)\}$
1	29	17	12	0.57	43	46	0.028
2	16	2	14	0.12	88	18	0.007
4	8	1	7	0.11	89	9	0.011
5	21	15	6	0.68	32	36	0.048
6	34	27	7	0.77	23	61	0.037
7	22	9	13	0.39	61	31	0.022
8	13	2	11	0.14	86	15	0.010
9	9	3	6	0.30	70	12	0.032
10	19	7	12	0.35	65	26	0.022
11	19	4	15	0.20	80	23	0.011
12	23	12	11	0.50	50	35	0.029
13	36	23	13	0.62	38	59	0.026
14	16	5	11	0.29	71	21	0.020
Sum	265	127				392	0.304
Avg.			10.6	0.39	61		0.023
Stdev.			3.1	0.22			

The calculations to determine $E(\theta)$ and $\text{VAR}(\theta)$ are shown in Table 6.47.

Table 6.47 : The extended 4-step procedure - calculations

STEP 1	STEP 2
$\lambda = 127$	$VAR(\lambda) = 127$
$\pi = 265$	$VAR(\pi) = 265$
STEP 3	STEP 4
$\delta = 10.6$	$VAR(\delta) = 3.1^2 = 9.6$
$\bar{\theta} = 0.39$	$VAR(\bar{\theta}) = 0.22^2 = 0.048$
STEP 5	
$E(\theta) = 0.39$	
$VAR(\theta) = 0.048 - 0.023 = 0.025 = 0.05^2$	
Avg(ARF) = 61 %	

Whereas the 'composite' analysis gave an overall ARF of 53 % in this an average ARF of 61 % has been calculated.

6.8 CONCLUSION

This Chapter presented both *Conventional* and *Bayesian* before-and-after methodologies to evaluate the effectiveness of road safety remedial measures at entity level as well as at group-of-entity level.

It was stated that any observed change in safety between the 'before' and 'after' periods after the application of a remedial measure could consist of 4 components - the *treatment* effect, the *regression-to-mean* effect, the *trend* effect and the *exposure* effect. *Conventional* and *Bayesian* methodologies were presented to account for the trend and exposure effects. It was concluded after, analysing the experimental data in Appendices A1 and A4, that the Bayesian method performs considerably better than the *Conventional* methods in eliminating the *regression-to-mean* effect even if short study periods are used. From the experimental studies it was found that at the group-of -entity level the length of the 'after' period had little influence on the magnitude of the estimated treatment effect.

A methodology was also presented to determine the effectiveness of a remedial measure at a location when there is reason to believe that the 'true' level of safety changed from year to year during both the 'before' and 'after' periods as a result of changes in traffic volumes and other confounding factors.

It was shown how each of all these methodologies can be reduced to a simple 4-step procedure to determine *measures of effectiveness* and their *variances*.

CHAPTER 7

MULTIVARIATE REGRESSION MODELS

7.1 INTRODUCTION

From Chapters 3, 4, 5 and 6 it is evident that multivariate regression models play an important role in applying the Empirical Bayesian methodology to estimate safety, to identify and rank hazardous locations and to evaluate the effectiveness of road safety remedial measures.

Multivariate regression models of data from a reference population are the source of the estimates $E(m)$ and $VAR(m)$ which serve as *prior* estimates for the Empirical Bayesian methodology.

An alternative method to estimate $E(m)$ and $VAR(m)$ is the *method-of-sample moments*. This method rests on the assumption that a linear relationship exists between the number of accidents and traffic flow at a location i.e. that the accident rate will remain constant irrespective of the traffic flow at a location.

Using multivariate regression models 'release' the Empirical Bayesian approach from this assumption. The regression model equation can be specified to allow for any hypothetical relationship between accidents and traffic flows.

It is often required to estimate the safety of an entity for which a sufficient reference population does not or cannot exist. Consider the example of a 2 lane rural road with lane widths of 3.55 m and 1.2 m gravel shoulders that has a traffic flow of 2345 vehicles per day. The ideal reference group, in order to apply the *method-of-sample moments*, would consist of a sufficient number of entities that share exactly the same geometric characteristics and traffic flows. It is unlikely that there will be many, or even any, entities that would meet these requirements.

By using a 'wider' reference group that consists perhaps of all 2-lane rural entities irrespective of lane widths, shoulder widths and traffic volumes a multivariate regression model can be developed from this reference group data to estimate precise values of $E(m)$ and $VAR(m)$ for any entity, irrespective of its geometric characteristics and traffic flows. For each entity the regression model serves to create, in the words of Hauer (1997), an 'imaginary' reference group.

Multivariate regression models allow for determining how the level of safety at an entity can expect to change with changing values of the covariates, such as for example, with changes in traffic flows. This ability of regression models forms the basis to predict what the level of safety would have been in the future had no road safety measures been undertaken.

7.2 MULTIPLE REGRESSION ANALYSIS – AN INTRODUCTION

According to Gujarati (1988) the modern interpretation of regression is as follows :

“Regression analysis is concerned with the study of the dependence of one variable, the dependent variable, on one or more other variables, the explanatory variables with a view to estimating and or predicting the population mean or average value of the former in terms of the known or fixed (in repeated sampling) values of the latter.”

The regression process begins with specifying the objectives of the regression analysis, including the selection of the dependent and independent variables. The next step is to design the regression analysis, considering such factors as sample size and the need for variable transformations. With the regression model formulated the assumptions underlying regression analysis are first tested for the individual variables. If all the assumptions are met then the model is estimated. Once results are obtained, diagnostic analyses are performed to ensure that the overall model meets the regression assumptions

and that no observations have undue influence on the results. The next stage is the interpretation of the modelling results. Finally the results are validated to ensure generalizability to the population. (Hair et al. ; 1995)

Probably the two most important steps of multivariate regression modelling is (Hauer ; 1997):

- a) The choice of model form (model equation).
- b) The estimation of parameters.

According to Hauer (1997) when developing a multivariate model these steps may have to be repeated several times. The results of one cycle of analysis provide motivation for modifications to the next cycle of analysis. Once parameters have been estimated and residuals examined, the form of the model may have to be revised, covariates may have to be added or dropped, and parameters estimated anew.

Hauer (1997) states that multivariate regression modelling is based on the belief that accident frequencies are associated in some orderly fashion with causal factors, and that this belief is embodied in a *model equation*.

A model equation states in what way accident frequencies are a function of the various covariate values

According to Hauer (1998) the choice of the model equation is more influential in determining the quality of the results than is the statistical technique to estimate the parameter values. The choice of model equation should reflect prior knowledge and beliefs about the nature of the relationship, and also insights obtained from an exploratory analysis of the data at hand. Hauer (1997) states that the choice of model equation reflects beliefs, embodies assumptions and impose limitations.

For example consider a model equation with the following form :

$$E(m_i) = L_i \alpha Q_i^\beta$$

$E(m_i)$ - *The expected accident frequency at location i .*

L_i - *The length of location i .*

Q_i - *The traffic flow at location i .*

The inclusion of L_i in the model equation embodies the belief that a road section along which all characteristics remain constant, will have a constant expected accident frequency per unit length along the whole length of the road. According to Hauer (1997) research has indicated that $E(m)$ may not be proportional to L if intersection accidents are included in the data set.

The inclusion of the α parameter serves the purpose of capturing the effect of all factors except for the effect of traffic flows. The influence of the traffic flow is accounted for separately by Q^β . The absence of a subscript i according to Hauer (1997) reflects the belief that influencing factors such as weather, economic conditions etc. have the same influence on all entities in the data set, and by implication, for the entire population which the data set is thought to represent.

The inclusion of Q^β reflects the belief that differences in Q make for differences in $E(m)$. The lack of a subscript i for the β parameter reflects the view that the safety of all entities depend on Q in the same manner. Using the form Q^β guarantees that $E(m) = 0$ as $Q = 0$. This form however has the limitation that it cannot represent the scenario where accidents initially increase with traffic flow, but at some point, as traffic flow increases further accident frequency begins to diminish.

Data from the reference group used in regression modelling serves to test the validity of the chosen model equation. If a covariate does not improve the fit of the model to the data consideration could be given to omitting that variable. If the introduction of a covariate improves the fit consideration could be given to

its inclusion in the model equation. Sometimes a model can be improved by transforming one or more variables, or even combining one or more variables.

It is important to note that a calibrated model only represents what statistical associations are present in the set of data at hand and that it does not necessarily capture plausible cause-effect relationships.

7.3 ORDINARY MULTIPLE LINEAR REGRESSION

7.3.1 MODEL FORM

The most common model form is one where the model is a linear function of its variables as well as its parameters.

$$Y = a_1 + a_2X_1 + a_3X_2 + \dots + a_{N+1}X_n + \varepsilon \quad \dots[7.2]$$

Y - *Level of safety e.g. accident rate or
= 1 for event
= 0 for non-event*

$a_1 - a_{n-1}$ - *Regression parameters*

$X_1 - X_n$ - *Independent variables e.g. roadway characteristics*

ε - *Error term*

To allow for a possible non-linear relationship between the dependent variable and an independent variable one or more polynomial terms are included in the model.

The need to include polynomial terms is generally indicated by means of an exploratory analysis of the general form of the relationship between a dependent variable and an independent variable. Constructing and then evaluating a scatter plot of these two variables is generally a common method.

To allow for the interaction of different driver, vehicle and/or roadway factors moderator (interaction) effects are often included in the model.

Dummy variables are also used where an independent variable is of a dichotomous or a categorical nature.

The following examples are presented to illustrate the different models that have been used to model accidents using linear regression analysis.

Bester (1988) developed the following model to predict the accident rate for all accidents using data from the Cape Province (now Western Cape, North West, Eastern Cape and Northern Cape) and the province of KwaZulu-Natal.

$$TAR = -0.295 - 0.12S^2 + 0.710T^2 + 0.933R_q - 0.648T(R_q) \quad \dots[7.3]$$

where $R^2 = 0.278$, and

S - *Shoulder width (m).*

T - *Topography : 1 = Flat ; 2 = Rolling and 3 = Mountainous.*

R_q - *Riding quality.*

The regression model contains two 2nd degree polynomials (S^2 and T^2) as well as an interaction term ($T \cdot R$).

The model shown above has a *additive* form. A linear regression model can also take a *multiplicative* form (Equation 7.4) which can then be linearised using a logarithmic transformation (Equation 7.5).

$$Y = a_1 X_1^{a_2} X_2^{a_3} \dots X_n^{a_{n+1}} \quad \dots[7.4]$$

$$\log(Y) = \log(a_1) + a_2 \log(X_1) + a_3 \log(X_2) + \dots + a_{n+1} \log(X_n) \quad \dots[7.5]$$

7.3.2 ASSUMPTIONS

According to Gebers (1998) the multiple linear regression method is based on the following assumptions :

- a) *Independence* - the Y observations are statistically independent of each other.
- b) *Linearity* – the value of Y is a linear function of $X_1, X_2, X_3 \dots X_n$
- c) *Homoscedasticity* – the variance of Y is the same for any fixed combination of $X_1, X_2, X_3 \dots X_n$.
- d) *Normality* – the errors of prediction (residuals) are normally distributed at all levels of Y.
- e) *Measurement infallibility* – the variates are free of measurement error.
- f) *Additivity* – the effect terms (coefficient X variable values) of the parameters can be combined in an additive fashion to estimate Y.

7.3.3 MODEL ESTIMATION

Linear multiple regression models are usually calibrated using the ordinary least square (OLS) method. The objective is to find the parameters that will minimise the following function (Gujarati ; 1988):

$$D = \sum_{i=1}^n (Y_i - \mu_i)^2 \quad \dots[7.6]$$

When the assumptions of homoscedasticity and/or independence do not hold, Gujarati (1988) recommends that the weighted least square method be used to estimate parameters.

7.3.4 PERFORMANCE

According to Vogt and Bared (1998) the use of linear regression models have produced disappointing results because of the following reasons :

- a) Accident frequencies and accident rates often do not follow a normal distribution.

Traffic accidents are random discrete events that are sporadic in nature. It is widely accepted that accident frequencies generally follow a Poisson distribution (Miaou et al. ; 1992). Normalising accident frequencies with exposure estimates to make accident rates appear to be a continuous random variable, according to Vogt and Bared (1992), do not change the fundamental discrete nature of accident data.

According to Gebers (1998) highly skewed Poisson-like variables produce heteroscedastic residuals, thereby introducing 'inconsistency' into the parameter estimates produced by OLS techniques. The use of OLS multiple regression in the presence of heteroscedasticity and non-normality results in regression models that, although not biased, do not satisfy the property of minimum variance.

- b) Accident frequencies for particular locations could be very small or even zero, even if several years of accident data have been obtained for those locations. Small integer counts, zero or close to zero do not typically follow a normal distribution. Using models calibrated using such small accident counts could predict negative values which is obviously not possible.

The presence of locations with zero accidents is particularly critical when using a *multiplicative* model linearized by a logarithmic transformation. The logarithm of zero is not defined. Locations with zero observations can therefore not be included in the investigation. Omitting such locations could be undesirable as traffic situations where

no accidents occur could be as important as locations where accidents occur.

Miaou and Lum (1993) concluded that linear regression models lack the distributional properties to adequately describe random, discrete, non-negative and sporadic vehicle accidents. Using linear regression method is therefore not a suitable method to estimate $E(m)$ and $VAR(m)$.

According to Miaou et al. (1992) the problems associated with the linear regression approach to modelling accidents can be overcome by using the Generalised Linear Modelling technique.

7.4 GENERALISED LINEAR MODELLING

According to McCullough and Nelder (1989) Generalised Linear Models (GLM) is a natural generalisation of classical linear models. The class of generalised linear models includes log-linear models for count data (such as accident counts).

The GLM approach allows the analyst to use dependent variables which may be bounded or which may have a non-continuous or a non-normal distribution. Accident frequencies for example are bounded in the sense that they can never be less than zero. As previously stated it also follows a non-normal distribution.

7.4.1 THEORETICAL FRAMEWORK

According to McCullough and Nelder (1989) the GLM approach assumes that there is a dependent variable Y and one or more independent variables (X_i 's) whose values influence the distribution of Y and that the independent variables influence the distribution of Y through a single linear function only. This function is called the *linear predictor* and is written as :

$$\eta = \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k \quad \dots[7.7]$$

The linear predictor may contain polynomials, moderator effects and/or dummy variables as long as it remains linear in its coefficients.

The linear prediction function is related to the expected value of an observation via a *link function*. There is one link function that is theoretically related to the error structure of the data on the basis of its underlying distribution. (Bonneson and McCoy ; 1997)

From McCullough and Nelder (1989) :

$$\eta = g(\mu_i) \quad \dots[7.8]$$

$$\mu_i = g^{-1}(\eta) = g^{-1}\left(\sum_{k=1} \beta_k X_{ik}\right) \quad \dots[7.9]$$

The GLM technique relies on the Maximum Likelihood Estimation approach to estimate the coefficients of the model. This approach involves estimating the β -coefficients that will maximise the log-likelihood function.

As with the link function there is a log-likelihood function that is theoretically related to the error structure of the data on the basis of its underlying distribution.

In the following section the application of the GLM approach to estimate log-linear Poisson and Negative Binomial models will be presented.

7.4.2 POISSON MODEL FORMULATION

According to Vogt and Bared (1998) the Poisson regression approach regards the number of accidents in a given space-time region as a random variable that takes values 0,1,2 with probabilities obeying the Poisson distribution.

Miaou et al. (1992) provides the following formulation of the Poisson regression model :

Consider a set of n locations. Let Y_i be the number of accidents of a particular type at location i during a year. Assume E_i is the exposure at this location. Associated with each location i there is a $k \times 1$ covariate vector denoted by $x_i = (x_{i1}, x_{i2}, \dots, x_{ik})'$, describing its geometric characteristics, traffic conditions and other relevant attributes. It is postulated that Y_i ($i = 1$ to n) are independent and each is Poisson distributed as :

$$p(Y_i = y_i | R_i = r_i, E_i = e_i, x_i) = \frac{(r_i e_i)^{y_i} e^{-(r_i e_i)}}{y_i!} \quad \dots[7.10]$$

Where

$$i = 1, 2, \dots, n$$

$$y_i = 0, 1, 2, \dots$$

R_i is the accident rate at location i and is expected to vary from one location to another depending on its covariates x_i . For each location i the Poisson model implies that the mean is equal to the variance.

If x_i and E_i are given with negligible uncertainties and R_i is assumed to be constant then Equation 7.10 becomes a classical Poisson regression model. The uncertainties in E_i and R_i introduce extra variations (or overdispersion) in the Poisson model. The consequences of ignoring the extra variations in the Poisson regression are that the maximum likelihood estimates (MLE's) of the regression coefficients, β , under the classical Poisson model are still consistent; however, variances of the estimated coefficients would tend to be underestimated. In other words the significance levels of the estimated coefficients may be overstated (Miaou et al. ;1992).

The link function for the Poisson model is (McCullough and Nelder ; 1989) :

$$\eta = \ln[\mu] \quad \dots[7.11]$$

$$\ln(\mu) = \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k \quad \dots[7.12]$$

$$\mu = \exp(\beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k) \quad \dots[7.13]$$

It is evident that the predicted value μ will always be greater than zero.

If the dependent variable is a rate which is expressed as accidents per unit of exposure then the expression above can be modified as follows :

$$\mu = E \cdot \exp(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k) \quad \dots[7.14]$$

$$\ln(\mu) - \ln(E) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k \quad \dots[7.15]$$

The logarithm of exposure is treated as an offset and is subtracted initially from the logarithm of the number of accidents before fitting the expression. This approach is based on the assumption that there is a linear relationship between exposure and accidents.

To allow for a possible non-linear relationship between accidents and exposure Equation 7.14 can be written as :

$$\mu = E^\theta \cdot \exp(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k) \quad \dots[7.16]$$

$$\ln(\mu) = \theta \ln(E) + \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k \quad \dots[7.17]$$

To estimate the safety of two-way stop controlled intersections on rural highways in Minnesota, USA, Bonneson and McCoy ((1993). developed a model with the following form :

$$A = \beta_0 T_m^{\beta_1} T_c^{\beta_2} \quad \dots[7.18]$$

A - Annual accident frequency.

T_m - Major road traffic demand (veh/day).

T_c - Minor road traffic demand (veh/day).

The *linear predictor* function for this type of multiplicative model takes the following form :

$$\eta = \ln(n) + \ln(\beta_0) + \beta_1 \ln(T_m) + \beta_2 \ln(T_c) \quad \dots[7.19]$$

Where n is the number of years of observations and $\ln(n)$ is termed the offset variable.

To predict the effect of median treatments on urban arterial safety Bonneson and McCoy (1997) used a model with the following form :

$$A = ADT^{\beta_1} L^{\beta_2} \exp(c_0 + c_1 X_1 + \dots + c_k X_k) \quad \dots[7.20]$$

where

A - Annual accident frequency.

ADT - Average daily traffic demand.

L - Street segment length.

X_i - Selected traffic and geometric characteristics.

In this case the *linear predictor* takes the form :

$$\eta = \beta_1 \ln(ADT) + \beta_2 \ln(L) + c_0 + c_1 X_1 + \dots + c_k X_k \quad \dots[7.21]$$

It is often found with the modelling of accidents that the variance of the dependant variable exceeds its mean. This is due to it not being practical nor possible to include and account for all variables that could influence accident occurrence. In such cases Vogt and Bared (1998) recommend that negative binomial models be used in accident modelling.

7.4.3 NEGATIVE BINOMIAL MODEL FORMULATION

According to Poch and Mannering (1996) Negative Binomial models generalise the Poisson form by permitting the variance to be overdispersed, equal to the mean plus a quadratic term in the mean whose coefficient is called the overdispersion parameter. When the parameter is zero a Poisson model results. When it is larger than zero, it represents variation above and beyond that due to the high way variables present in the model.

If accident frequencies are overdispersed the probability density function of y_i is as follows (Vogt and Bared ; 1998):

Where k is the over-dispersion parameter.

$$P(y_i) = \frac{\Gamma(y_i + k)}{y_i! \Gamma(k)} \left(\frac{k^{-1} \mu_i}{1 + k^{-1} \mu_i} \right)^{y_i} \left(\frac{1}{1 + k^{-1} \mu_i} \right)^k \quad \dots[7.22]$$

And

$$\text{var}(y_i) = \mu_i + \frac{(\mu_i)^2}{k} \quad \dots[7.23]$$

If $K = 1/k = 0$ then the negative binomial reduces to the Poisson model. The larger the value of K the more variability there is in the data over and above that associated with the mean μ .

The link function for the Negative Binomial model is (Bonneson and McCoy ; 1993):

$$\eta = \ln \left[\frac{\mu}{k + \mu} \right] \quad \dots[7.24]$$

The different linear predictors presented for the Poisson model can also be used for the Negative Binomial model.

7.4.4 LOG LINEAR MODEL ESTIMATION

The GLM technique uses Maximum Likelihood Estimation (MLE) methods to calibrate the regression models. This method entails finding the β -coefficients and in the case of the Negative Binomial also the k coefficient that will maximise the following log-likelihood functions :

Poisson Model

$$\ln(L) = \sum_{i=1}^n y_i \log(\mu_i) - \mu_i \quad \dots[7.25]$$

(from McCullough and Nelder ; 1989)

Negative Binomial Model

$$\ln(L) = \sum_{i=1}^n y_i \log(ky_i) - (y_i + k^{-1}) \log(1 + k\mu_i) + \log \left(\frac{\Gamma(y_i + k^{-1})}{\Gamma(y_i + 1)\Gamma(k^{-1})} \right) \quad \dots[7.26]$$

(from SAS Institute Inc. ; 1999)

The regression parameters can be determined by maximising the log-likelihood function by an iterative weighted least square procedure. (McCullough and Nelder ; 1989).

According to McCullough and Nelder (1989) an equivalent approach to maximising the log-likelihood function is by minimising the *deviance*.

The deviance is given by :

$$D = 2(L^f - L^m) \quad \dots[7.27]$$

Where

L^f - *Log-likelihood value for the 'null' model.*

L^m - *Log-likelihood value for the model under consideration.*

The log-likelihood value for the 'null' model can be determined from fitting a model to the accident data that consist of a constant term only (no variables included).

According to McCullough and Nelder (1989) the deviance function for a Poisson model is :

$$D = 2 \sum (Y_i \ln(Y_i / \mu_i) - (Y_i - \mu_i)) \quad \dots[7.28]$$

Gebers (1998) states that a fundamental assumption underlying ordinary least squares linear regression analysis is that all random errors have the same variance at difference values of the explanatory variable. The homogeneity of residual error assumption is invariably violated with accident data because of the direct proportional relationship between the means and variances of the arrays, thereby introducing heteroscedasticity into the distribution of the residuals. For this reason the OLS technique is not suitable to calibrate models based on the GLM technique.

Saccomano and Buyco (1988) demonstrated that the weighted least squares algorithm (WLSA) for calibrating log-linear models produces high residuals for cells that are characterised by low cell memberships in the contingency table

of causal factors. The WLSA approach requires large samples. Koch and Imrey (in Saccomanno and Buyco ; 1988) noted that the WLSA approach is sensitive to small observed and expected cell counts. This approach is especially problematic for assessing situations where there is exposure but no accidents – as could easily be the case with low volume roads. The MLE technique can however accommodate such a scenario.

7.4.5 ESTIMATION TECHNIQUES

A common approach to estimate the coefficients of log-linear models is to use an iterative procedure. Bonneson and McCoy (1992). describe the procedure as follows :

First the data is analysed using a Poisson model. Secondly, a regression is fitted to Eqn. 7.30, using MLE techniques to estimate the value of K and a measure of its degree of significance.

$$e_i^2 = \mu_i + \frac{\mu_i^2}{k} \quad \dots[7.30]$$

where

$$e_i^2 = (Y_i - \mu_i)^2 \quad \dots[7.31]$$

The need for a 3rd analysis step is based on the dispersion parameter (σ_d) and the significance of the k parameter (from Step 2), where :

$$\sigma_d = \frac{D^m}{n - p} \quad \dots[7.32]$$

D^m - The deviance of the Poisson model (From Step 1).

n - Total number of observations.

p - Total number of estimated coefficients.

If the dispersion parameter is greater than 1 and the k parameter is significant, a 3rd analysis step is conducted using the Negative Binomial model with the value of k from Step 2. The residuals from this analysis are then analysed in a 4th step to determine a new k parameter. Steps 3 and 4 are repeated until convergence on a value of k .

7.4.6 ESTIMATION TOOLS

Generalised linear models can be estimated using specialised statistical software. The most common software packages used by road safety researchers to estimate log-linear models are the GLIM software package and the non-linear regression procedure (NLIN) in the SAS statistical software.

7.4.7 MODEL EVALUATION

According to Vogt and Bared (1998) there are three important tests for an acceptable model. These are as follows:

- a) The estimated regression coefficient for each independent variable should be statistically significant i.e. one should be able to reject the null hypothesis that the coefficient is zero.
- b) Engineering and intuitive judgements should be able to confirm the validity and practicality of the sign and rough magnitude of each estimated coefficient; and
- c) Goodness-of-fit measures and statistics should indicate that the variables do have explanatory and predictive power.

An important assumption of the GLM technique is that there should be no correlation between different observations. In other words all the observations should be independent of one another. There are situations in which a possibly of dependence between observations may arise, for instance treating

each year of accident data at a particular site as a separate observation, or considering the accidents on intersection approaches as separate observations, where it is quite possible that accident numbers on the approaches of the same intersection might be correlated with one another.

McCullough and Nelder (1989) propose that goodness-of-fit be evaluated on the basis of the deviance or on the generalised Pearson X^2 statistic.

The deviance of a model m is :

$$D^m = 2(L^f - L^m) \quad \dots[7.32]$$

Where L^f is the log-likelihood that would be achieved if the model gave a perfect fit and L^m is the log-likelihood of the model under consideration.

The overdispersion parameter (σ_d) is given by :

$$\sigma_d = \frac{D^m}{n - p} \quad \dots[7.33]$$

where

D^m - Deviance.

n - Number of observations.

p - Number of coefficients estimated.

According to Vogt and Bared (1998) if the value of the overdispersion parameter (σ_d) is larger than 1 then the Poisson model is inappropriate because of overdispersion due to missing variables and/or measurement error. In such cases a Negative Binomial models is generally indicated.

The Pearson chi-square statistic is defined by (McCullough and Nelder ; 1989) as:

$$X^2 = \sum_{i=1}^n \frac{(y_i - \mu_i)^2}{V(y)} \quad \dots[7.34]$$

Where $V(y)$ refers to the variance function of the model.

The variance functions for the Poisson and the Negative Binomial models are as follows:

Poisson : $V(y) = \mu_i \quad \dots[7.35]$

Negative Binomial : $V(y) = \mu_i + (\mu_i^2)/k \quad \dots[7.36]$

For a valid Poisson model the *Pearson* X^2 statistic follows a chi-square distribution with $n-p$ degrees of freedom. Therefore if $X^2/(n-p)$ equals approximately 1 the assumed Poisson error structure is approximately equivalent to that found in the data. Should the statistic be significantly larger than 1 then overdispersion is indicated and the Negative Binomial model should be considered.

For a Negative Binomial model the *Pearson* X^2 statistic has a chi-square distribution with $n-p-1$ degrees of freedom. Thus if $X^2/(n-p-1)$ equals approximately 1 the assumed Negative Binomial error structure is approximately equivalent to that found in the data. (Bonneson and McCoy ; 1992).

According to McCullough and Nelder (1989) the *Pearson* X^2 statistic is not well defined in terms of minimum sample size when applied to non-normal distributions. They recommend that the *Pearson* X^2 statistic should not be used as an absolute measure of model significance.

The R^2 (*coefficient of determination*) goodness-of-fit measure, used to estimate the percentage of variation explained by a linear regression model is not suitable for Poisson and Negative Binomial regression models. In some

instances the R-square measure may provide values greater than one, particularly for non-linear models. (Vogt and Bared ; 1998)

A more subjective measure of model fit can be obtained from a graphical plot of prediction ratio vs. the estimate of the expected accident frequency. The prediction ration is described as the normalised residual (Bonneson and McCoy ; 1992) :

$$PR_i = \frac{\mu_i - Y_i}{\sqrt{V(y)}} \quad \dots[7.38]$$

where $V(y)$ is the variance function.

This type of plot yields a visual assessment of the predictive ability of the model over the full range of μ_i 's. A well fitting model would have the prediction ratios symmetrically around zero.

7.5 SUMMARY and CONCLUSION

This Chapter provided a brief overview of the processes involved in developing accident prediction models using multivariate regression techniques. It was shown that the classical linear regression approach is not a suitable approach to model accident frequencies and accident rates. An alternative approach called Generalised Linear Modelling (GLM) was presented and details on two log linear modelling approaches – Poisson regression and Negative Binomial were presented. It was shown that these modelling approaches are suitable to model accident data. To illustrate how the concepts discussed are applied in practice a case study will be presented to show how the Negative Binomial modelling approach was used to estimate the safety of signalised intersections.

CASE STUDY 7.1

ESTIMATION OF SAFETY AT SIGNALISED INTERSECTIONS

By Hauer E, Ng JCN and Lovell J (1988)

Introduction

Hauer et al. (1988) developed a number of microscopic and macroscopic models to measure the safety at signalised intersections if the vehicular flows using it are known.

Data

They selected for analysis a set of intersections that are similar in most respects except for traffic flows and accident history. The data set consisted of accident data and traffic flow information for 145 four-legged, fixed-time, signalised intersections in Metropolitan Toronto that carry two-way traffic on all approaches and have no turn restrictions. Most are on straight, level sites with a speed limit of 60 km/h.

The accident data was collected over a three year period – 1982 to 1984. For each approach one day manual traffic counts were conducted. Turning and through moving flows were thus available for each approach for the AM-peak, the PM-peak and the off-peak periods. All vehicle counts are for weekday conditions.

The analysis was confined to the AM peak (07:00 – 09:00) the PM-peak (16:00 – 18:00) and the following off-peak period - 10:00 to 15:00. The periods from 09:00 to 10:00 and 15:00 to 16:00 were excluded because the signal timings changed during these periods.

The accident dataset only contained information on accidents involving two vehicles. This type of accident accounted for 81 % of the total accidents.

Model development

One of the objectives of the study was to relate accidents to the traffic flows to which they belong. Fifteen (15) accident patterns were identified and are shown in Figure 6. Accidents in each pattern were defined by the manoeuvres of the two vehicles before colliding. The common categorisation by initial impact type such as rear-end, angle, turning movement, sideswipe etc. was avoided because of its ambiguity.

The next step was to decide on the model form for each one of these 15 accident patterns. The guiding principle in this process was the wish to ensure a satisfactory fit with parsimony of parameters and without violation of the obvious logical requirements.

To determine the most appropriate model form exploratory data analysis was first undertaken. The following example is presented on how a model form was developed for Accident Pattern 6.

Table 7.1 gives the average number of accidents per site in 3 years for five ranges left¹ turning flows versus five ranges of the straight-through flow. The irregular flow ranges were selected so that each row and column would have approximately one-fifth of all accidents

¹ 'Left' refers to the Canadian/USA driving convention. This should read 'right' for a South African interpretation.

Table 7.1 : Traffic flows for accident pattern 6.

Straight-Through Flow	Left turning flow					ROW TOTAL
	0-521	522-821	822-1033	1034 - 1408	1409 – 4038	
0 - 3525	0.49	0.60	0.63	1.57	0.56	0.64
3526 - 4825	0.61	1.00	1.91	1.24	1.75	1.16
4826 - 5941	1.27	1.22	1.44	1.36	2.79	1.63
5942 - 7771	1.31	1.40	3.10	2.20	1.40	1.78
7772 - 12091	1.67	2.75	2.63	4.50	3.27	2.73
Total	0.74	1.05	1.56	1.93	1.91	1.26

The relationship between accidents and left-turning flows are shown in Figure 7.1.

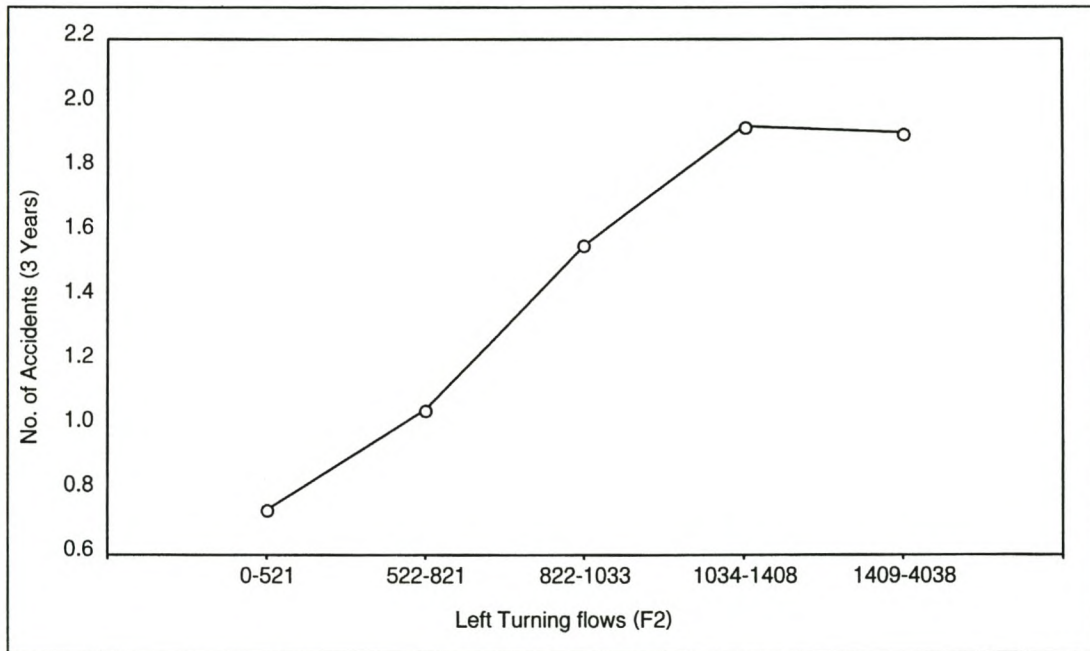


Figure 7.1 : Accident frequency vs. left turning flows.

The relationship between straight through flows and accident frequency is shown in Figure 7.2.

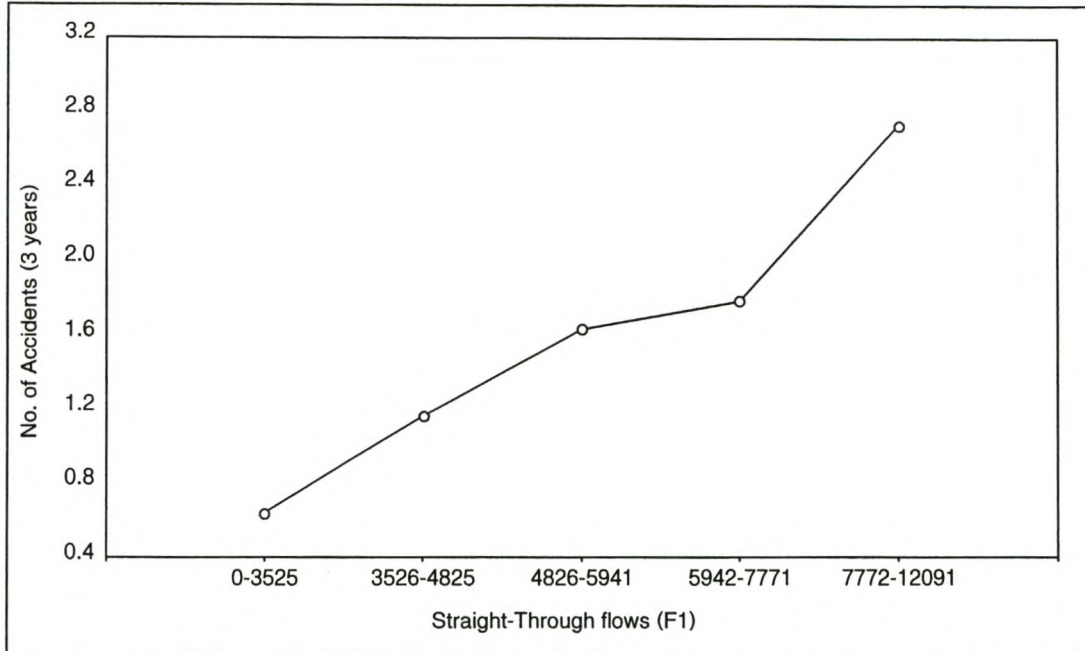


Figure 7.2 : Accident frequency vs. straight through flows

From Figure 7.2 it appears that accidents are proportional to the through traffic (F_1) i.e. $A \propto F_1$ and from Figure 7.1 that the increase of accidents with the left turning flow (F_2) appears to be non-linear i.e. $A \propto F_2^{b_1}$ where $b_1 < 1$.

A suitable model form for accident pattern 6 is thus :

$$A = aF_1F_2^{b_1}$$

This process of data analysis was applied to all 15 accident patterns and the subsequent model forms are indicated in Table 7.2.

All these models are *mutiplicative* in nature and it was assumed they all have a negative binomial error structure.

Coefficient Estimation

The coefficients were estimated using the GLIM (Generalised Linear Interactive Model). The process of coefficient estimation was iterative. Firstly a value for k was assumed and then proceeded to estimate the β -coefficients using GLIM. The residuals were calculated and from the residuals a new value of k was determined. This k value was then fed back into the model to calculate a new set of β -coefficients. The residuals and a new k value were calculated again and this process was repeated until convergence.

Interpretation

Hauer et al. (1998) concluded that a closer examination of how the frequency of accidents depend on traffic flows from which they arise reveals that preconceived ideas are sometimes not supported by empirical evidence. The results in Table 7.2 reveal that accident frequencies between vehicles travelling in the same direction is proportional to the traffic flows in that direction (Patterns 1 and 2). The frequency of accidents between left turning and through moving vehicles is proportional to the flow of through traffic but less than proportional to the flow of left-turning vehicles. (Pattern 6). The frequency of accidents between right-angle flows (Pattern 4) is not influenced by the major road flow but by the minor road flow which exerts a great deal of influence initially, but that this tapers off as flows become larger.

According to Hauer et al. (1988) these observations lead to the conclusion that the popular assumption that intersection accidents are proportional to the sum of entering volumes is not supported by empirical evidence for common accident types. They state it is therefore not correct to use intersection accident rates calculated on the basis of the sum of entering volumes to compare the safety of two different intersections, or between 'before' and 'after' periods.

Table 7.2 : Results of multivariate regression analysis

Pattern	Model Form	Time	a	b	c	k
1	A = a(F)	AM	$0.1655 * 10^{-6}$			2.98
		PM	$0.2178 * 10^{-6}$			2.73
		Off	$0.2164 * 10^{-6}$			3.54
		daily	$0.2052 * 10^{-6}$			4.59
2	A = a(F)	AM	$0.0987 * 10^{-6}$			1.49
		PM	$0.0933 * 10^{-6}$			0.94
		Off	$0.1080 * 10^{-6}$			4.15
		daily	$0.1014 * 10^{-6}$			1.97
3	A = a(F ₂) ^c	Daily	$8.6129 * 10^{-9}$		1.0682	1.20
4	A = a(F ₂) ^c	AM	$19.020 * 10^{-6}$		0.1536	2.65
		PM	$1.4127 * 10^{-6}$		0.6044	2.33
		Off	$9.7329 * 10^{-6}$		0.3860	3.38
		Daily	$8.1296 * 10^{-6}$		0.3662	5.51
5	A = a(F ₁) ^b (F ₂) ^c	Daily	$0.3449 * 10^{-6}$	0.1363	0.6013	1.2
6	A = a(F ₁)(F ₂) ^c	Daily	$0.0418 * 10^{-6}$		0.4634	2.1
7	A = a(F ₁) ^b (F ₂) ^c	Daily	$0.2113 * 10^{-6}$	0.3468	0.4051	1.2
8	A = a(F ₂) ^c	Daily	$2.6792 * 10^{-6}$		0.2476	1.2
9	A = a(F ₁) ^b	Daily	$6.9815 * 10^{-9}$	1.4892		1.2
10	A = a(F ₂) ^c	Daily	$5.590 * 10^{-12}$		2.7862	1.2
11	A = a(F ₁) ^b (F ₂) ^c	Daily	$1.3012 * 10^{-9}$	1.1432	0.4353	1.2
12	A = a(F ₁) ^b (F ₂) ^c	Daily	$0.0106 * 10^{-6}$	0.6135	0.7858	1.2
13	A = a(F ₁) ^b (F ₂) ^c	Daily	$0.4846 * 10^{-6}$	0.2769	0.4479	1.2
14	A = a(F ₁) ^b (F ₂) ^c	Daily	$1.7741 * 10^{-9}$	1.1121	0.5467	1.2
15	A = a(F ₁) ^b	Daily	$0.5255 * 10^{-6}$	0.4610		1.2

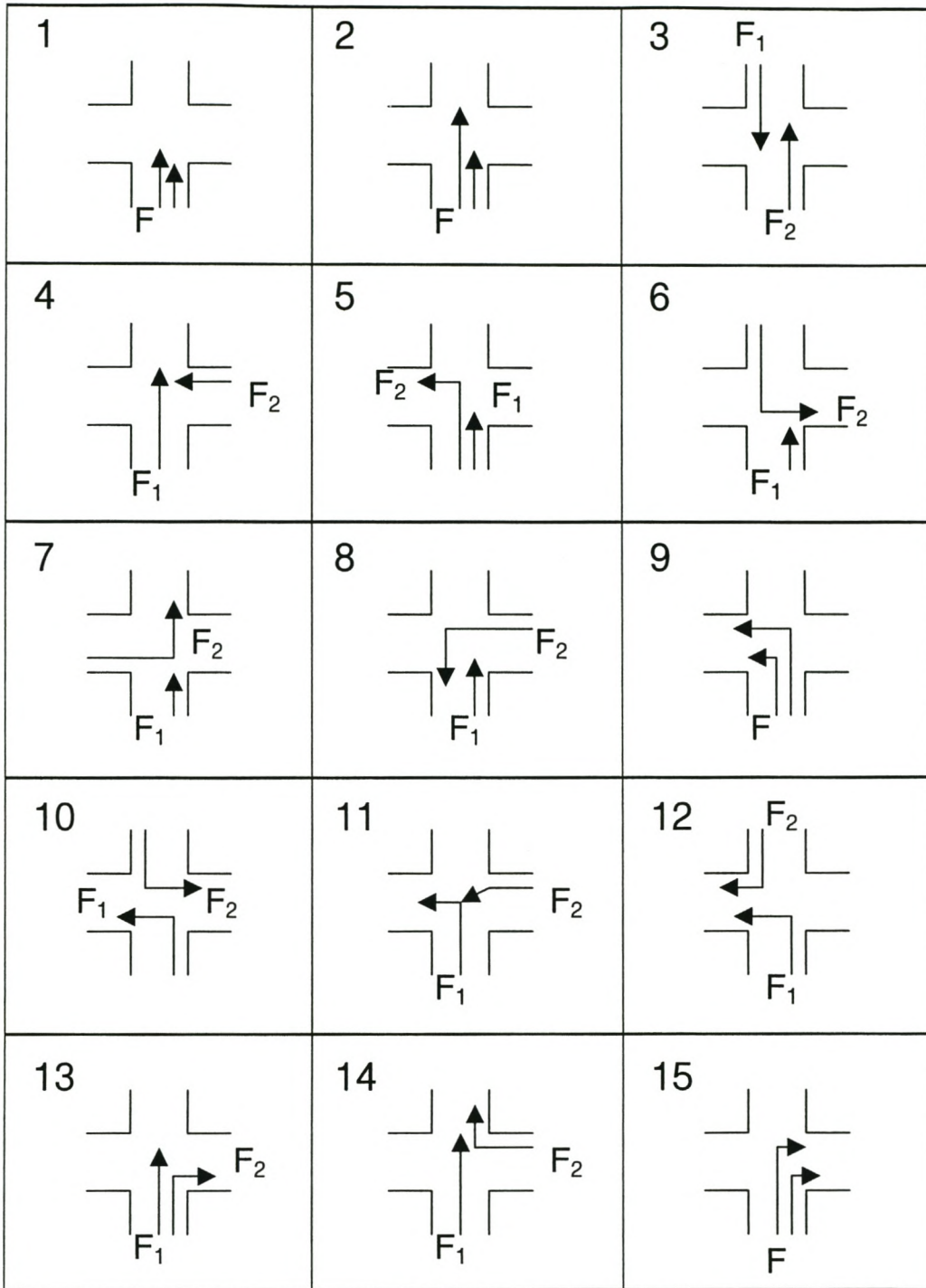


Figure 7.3 : Intersection accident patterns (Right hand driving rule)

CHAPTER 8

SUMMARY AND CONCLUSIONS

When implementing road safety remedial programmes it is required to analyse road traffic accident data in a manner that will ensure the effective and efficient utilisation of often scarce resources.

This thesis has shown how accident data could be analysed to measure safety at road entities, to identify and rank hazardous locations and how to determine the effectiveness of road safety remedial measures.

Having valid and reliable estimates of the true level of safety of an entity or a group of entities is fundamentally important for the efficient identification and ranking of hazardous locations and determining the effectiveness of remedial measures.

The methods available to measure safety can be divided into two categories, a) Conventional methods and b) Bayesian methods. The Conventional methods are based on the assumption that the level of safety at a location is fixed and that it can be determined by the observed accident experience, while the Bayesian methods assume that the true level of safety at a location is unknown, a variable and that the observed accident experience can provide a clue as to what the real value is.

It has been shown that the conventional accident rate measure to measure safety on road segments produces results that are generally not reliable. The main reason for this is that a non-linear relationship normally exists between accident frequencies and exposures. It can be concluded therefore that should the Safety Performance Function be non-linear accident rates should not be used to compare time periods nor different locations with each other. Similarly it has been shown that the common *sum-of-flows* accident rate measure to measure the safety of intersections is logically unsatisfactory for a

number of reasons. Using a safety measure based on the product of flows is generally considered to be more appropriate.

Safety Performance Functions can be used to overcome the inherent shortcomings of the conventional accident rate measures for segments and intersections. However SPF's are prone to a certain degree of bias if average traffic flow values such as AADT's are used, or if the SPF serves as an aggregate function for different accident types, each of which are associated with significantly different levels of risk. Depending on the form of the SPF, the chosen accident measure, the variability in traffic flows etc. this bias could be significantly large.

When identifying hazardous locations it has been shown that it is important to use methodologies that will maximise the number of *true positive* identifications and minimise the number of *false positive* identifications. As illustrated in Example 5.11 it could be the case that a large degree of *true positive* identification can only be obtained at the expense of the *false positive* identification rate. In assessing the overall efficiency of a hazardous location identification method it becomes necessary to weigh the relative cost and consequences of *false negative* and *false positive* identifications. A *false negative* identification would let a true hazardous location continue untreated with potentially very serious consequences in terms of loss of life, injury and damage to property, while a *false positive* identification could cause scarce resources to be allocated to locations with a low potential for accident reduction.

The identification methods presented in this work were divided into two categories, namely : a) Conventional methods, and b) Bayesian methods. Applying these methods to the experimental data in Appendix A1 revealed that it appears the Bayesian methods could produce lower *true positive* identification rates as well as lower *false positive* identification rates than the Conventional methods.

It could often be the case that to further investigate all the identified hazardous locations would require more financial resources than are available. Decisions are then required on which hazardous locations to investigate and which not to investigate further. The most common approach is to rank the identified hazardous locations according to some criteria and then to start investigating from the 'top' in descending order until the budget for such road investigations are depleted.

A number of possible criteria to rank hazardous locations were presented. These criteria can be roughly divided into three groups – those that make use of *Conventional* safety estimates, and those make use of *Bayesian* safety estimates and regression models, and finally those that combine *Conventional* or *Bayesian* estimates with accident severity information.

Criteria based on accident rates have been shown to produce a bias in favour of locations with low accident totals and low traffic volumes i.e. sites whose accident reduction and potential economic returns are low. Criteria based on accident numbers have been shown to produce a bias in favour of locations with high accident totals and traffic volumes. Since one of the overall objectives of a remedial programme is to improve the efficiency of the road transportation system by reducing accidents, the PAR (Potential Accident Reduction) index is a popular method of ranking hazardous locations. The PAR index provides a measure of the expected accident reduction at a location and therefore also a measure of the expected economic return at a location.

To quantify the potential accident reduction associated with specific remedial measures (e.g. shoulder widening) the PAR index can be estimated using the Bayesian safety estimates and appropriate multivariate regression models. Research by Persaud et al. (1999b) has shown that this approach generally produces better results than methods based solely on *Conventional* and *Bayesian* safety estimates. Persaud et al. (1999a) have also shown how this method of estimating the PAR index can be combined with accident severity information and accident costs to rank hazardous locations.

Evaluating the effectiveness of a remedial measure/s is a very important last step when implementing an accident remedial programme. This knowledge is required to determine the actual cost effectiveness of a remedial measure and to provide factual guidance for the future.

The determination of the effectiveness of a remedial measure basically involves estimating what would have been the level of safety in the 'after' period had the remedial measure not been implemented and then to compare this estimate with the level of safety actually observed in the 'after' period.

The observed number of accidents in the 'before' period are most often used in procedures to determine the '*what would have been*' level of safety in the 'after' period. In doing this a number of factors should be considered and allowed for, these are – a) the *regression-to-mean* effect, b) *exposure* effect and c) *trend* effect.

The *regression-to-mean* effect occurs because of the non-random manner in which hazardous locations are chosen. Since hazardous locations are normally selected on the basis of a high number of accidents or accident rate in the before period it is reasonable to expect that, because of the random nature of accidents, the accident rate/number in the after period could be less even in the absence of any remedial measures. The *regression-to-mean* effect is particularly relevant when *conventional safety* estimates are used. It has been shown however that the magnitude of this effect decreases as the study period increases.

The *regression-to-mean* effect can largely be eliminated by using Bayesian safety estimates.

The *exposure* effect becomes particularly relevant if there has been an increase/decrease in the traffic volumes using a location between the 'before' and 'after' periods. *Conventional* and *Bayesian* methodologies were presented

on how to eliminate the exposure effect when determining the effectiveness of a remedial measure.

The *trend* effect is of importance if there have been changes in the traffic, road, social environments etc. between the 'before' and 'after' periods. Changes in the traffic composition, climate, urban structure, law enforcement quality/quantity etc. could all for example have an impact on the level of safety at a location. The *trend* effect is accounted for by using a *comparison group*. A methodology is presented on how to choose the most appropriate comparison group for a *before-and-after with comparison group* study. Guidelines have also been provided on how to design such a study and how to interpret the results.

A Bayesian methodology based on the use of multivariate regression models for each year in the study period is a suitable methodology to eliminate the *regression-to-mean* effect and also to account for any *trend* and *exposure* effects. This methodology as proposed by Hauer (1997) is an improvement on other Bayesian and Conventional methodologies because it does not assume that the 'true' level of safety during either the before or after periods remained constant for the whole period. It is therefore based on the realistic assumption that the true level of safety of a location may change from year to year for a variety of reasons.

Many of the methodologies to measure safety, to identify and rank hazardous locations and to determine the effectiveness of remedial measures rely on the availability of suitable multivariate regression models. It has been shown that the ordinary least squares regression technique produce models that are not suitable. Instead the generalised linear modelling approach was presented as an alternative.

CHAPTER 9

RECOMMENDATIONS

It is recommended that road authorities use the Bayesian methodologies in this thesis to analyse accident data when implementing road safety remedial programmes. Before this can be undertaken it is recommended that the following critical issues be addressed :

- a) Clarify and/or improve the legal definition of what constitutes an accident and ensure that all reporting officers interpret and apply the definition consistently.
- b) Improve the management of accident data, particularly focussing on eliminating underreporting, missing data and incorrect/incomplete locational information.
- c) To ensure that all fatalities that occur within 6 days are recorded as such and that follow-up procedures for casualties admitted to hospital should be improved. Attention could be given to linking hospital databases and accident databases.
- d) Bayesian accident data analysis methodologies should be included in the graduate and/or post graduate civil engineering curricula.
- e) Dedicated research should be undertaken by research institutions to develop appropriate Safety Performance Functions (accident models) for different road and intersection types.
- f) Software systems should be developed to assist local authorities in implementing the Bayesian methodologies in this thesis.

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APPENDIX A1

EXPERIMENTAL DATA

- **Introduction**

In order to compare different measurement, identification, ranking and evaluation methods, and also to assess the influence of different study periods on the performance of these methods a number of experimental studies were performed.

These studies were performed using data generated randomly for a total of 1000 sites. The true level of safety at each site (m_i) was randomly generated from a Gamma distribution with mean = 4 acc/year and variance = 2. Based on the assumption that at each site the observed annual accident frequencies will follow a Poisson distribution around the true mean (m_i), a total of 5 years of accident frequency data were randomly generated.

The moments of the Gamma distribution were chosen arbitrary. It is assumed that in practice it is likely that there exist a target accident definition and a reference group of sites that will follow a similar Gamma distribution. Only 5 years of accident frequency data were generated because in practice more than 5 years worth of data are seldomly used. To keep the experiment simple it was assumed that traffic volumes and other possible confounding factors remained absolutely constant over the 5 year period.

- **Procedure**

The experimental data was generated using a spreadsheet application – Microsoft® Excel®.

Below the random numbers, the true level of safety (m_i) and the annual accident frequencies for the first three sites are shown :

	A	B	C	D	E	F	G	H	I	J	K	L	M
1	ID	Rnd1	Rnd2	Rnd3	Rnd4	Rnd5	Rnd6	m_t	X_1	X_2	X_3	X_4	X_5
2	1	0.323	0.251	0.098	0.642	0.716	0.931	3.24	2	1	4	4	6
3	2	0.301	0.944	0.524	0.315	0.850	0.280	3.16	6	3	2	5	2
4	3	0.912	0.573	0.381	0.721	0.063	0.347	6.01	6	5	7	3	5

The following procedure was followed :

- Using the `=rand()` function of Microsoft® Excel® for each of the 1000 sites 6 random numbers were generated – *rnd1*, *rnd2*, *rnd3*, *rnd4*, *rnd5* and *rnd6*.
- Random number *Rnd1* was used to generate the true level of safety at a site (m_t) assuming it is from a Gamma distribution with mean = 4 acc/year and variance = 2. The Gamma parameters $\beta = 8$ and $\alpha = 2$. The `=GAMMAINV` function of Excel® was used to generate m_t as follows-

$$m_t = \text{GAMMAINV}(\text{rnd1}, 8, 0.5)$$

In Excel *alpha* = β and *beta* = $1/\alpha$.

- Using the Microsoft Excel macro below random number *Rnd2* was used to generate X_1 , *Rnd3* to generate X_2 , *Rnd4* to generate X_3 , *Rnd5* to generate X_4 and *Rnd6* to generate X_5 . (Refer to Table above)

Sub poisson()

Dim i As Integer

Dim j As Integer

Dim f As Double

Dim fo As Double

Dim fs As Double

Dim dum As Integer

```

For k = 0 to 4
  For i = 0 to 999
    fs = 0
    j = 0
    while fs <= Cells(i + 2, 1 + k)
      If j = 0 Then fo = Exp(-1 * (Cells(i + 2, 8)))
      If j > 0 Then fo = fo * (Cells(i + 2, 8) / j)
      fs = fs + fo
      Cells(i + 2, 9 + k) = j
      j = j + 1
    Wend
  Next
Next
End Sub

```

The objective of this macro is the estimate the value of j at which :

$$\sum_{j=0}^j \frac{m_t^j e^{-m_t}}{j!} \leq Rnd$$

- **Data summary**

Table A1.1 shows summary statistics for m_t , X_1 , X_2 , X_3 , X_4 and X_5 .

Table A1.1 : Summary Statistics

Statistic	m_t	X_1	X_2	X_3	X_4	X_5
Mean	4.025	4.047	3.989	3.985	4.059	4.114
Median	3.859	4	4	4	4	4
Mode	-	3	4	3	3	3
Mode frequency	-	164	170	174	173	204
Minimum	0.591	0	0	0	0	0
Maximum	10.12	12	16	18	15	13
25th Percentile	2.958	2	2	2	2	2
75th Percentile	4.887	5	5	5	6	5
Std.dev.	1.408	2.385	2.577	2.476	2.525	2.437
Variance	1.982	5.688	6.640	6.131	6.376	5.941
Sum	4025	4047	3989	3985	4059	4114
Dispersion (k)	-	9.978	6.003	7.400	7.111	9.264

According to Vogt and Bared (1998) if m_t has a Gamma distribution between sites and X is Poisson distributed about m_t then the distribution of X between sites would follow the Negative Binomial distribution, where :

$$E(X) = \text{Mean and } \text{VAR}(X) = \text{Mean} + (\text{Mean})^2/k$$

Figure A1.1 shows the fitting of a Gamma distribution to the generated values of m_t .

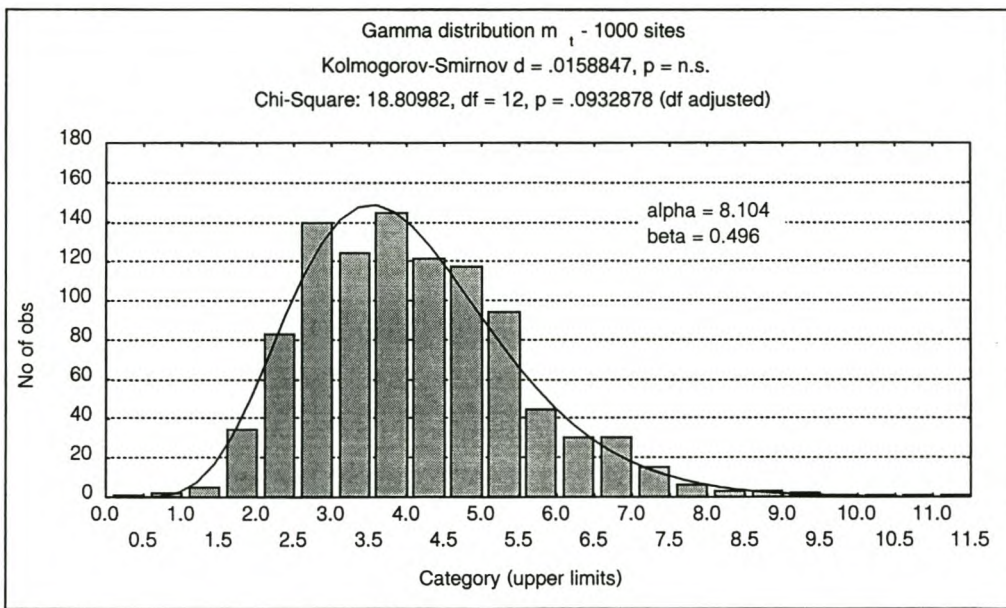


Figure A1.1 : Fitting of Gamma distribution

- **Data listing**

A complete list of the random numbers used and the generated values for m_t , X_1 , X_2 , X_3 , X_4 and X_5 for all 1000 sites are shown in Table A2.1

Table A1.2 : Listing of data for experimental sites (Sites 1 to 50)

ID	rnd1	rnd2	rnd3	rnd4	rnd5	rnd6	m_t	X_1	X_2	X_3	X_4	X_5
1	0.323	0.251	0.098	0.642	0.716	0.931	3.24	2	1	4	4	6
2	0.301	0.944	0.524	0.315	0.850	0.280	3.16	6	3	2	5	2
3	0.912	0.573	0.381	0.721	0.063	0.347	6.01	6	5	7	3	5
4	0.377	0.766	0.376	0.369	0.755	0.982	3.42	5	3	3	5	8
5	0.631	0.787	0.609	0.711	0.628	0.736	4.31	6	5	5	5	6
6	0.191	0.725	0.455	0.926	0.548	0.701	2.75	4	2	5	3	3
7	0.137	0.377	0.775	0.252	0.110	0.785	2.52	2	4	1	1	4
8	0.504	0.955	0.720	0.388	0.164	0.349	3.85	7	5	3	2	3
9	0.255	0.716	0.571	0.576	0.204	0.832	3.00	4	3	3	2	5
10	0.073	0.617	0.777	0.630	0.437	0.018	2.16	2	3	2	2	0
11	0.054	0.783	0.053	0.721	0.741	0.733	2.03	3	0	3	3	3
12	0.156	0.613	0.466	0.226	0.765	0.567	2.60	3	2	1	4	3
13	0.162	0.889	0.601	0.448	0.260	0.854	2.63	5	3	2	1	4
14	0.605	0.010	0.632	0.484	0.228	0.971	4.21	0	5	4	3	8
15	0.602	0.286	0.528	0.761	0.710	0.204	4.20	3	4	6	5	2
16	0.063	0.641	0.703	0.722	0.476	0.327	2.09	2	3	3	2	1
17	0.193	0.121	0.390	0.283	0.551	0.714	2.76	1	2	2	3	4
18	0.166	0.567	0.608	0.578	0.460	0.899	2.65	3	3	3	2	5
19	0.797	0.046	0.179	0.281	0.238	0.504	5.10	2	3	4	3	5
20	0.807	0.105	0.112	0.670	0.939	0.386	5.16	2	2	6	9	4
21	0.428	0.872	0.450	0.780	0.014	0.773	3.59	6	3	5	0	5
22	0.186	1.000	0.007	0.194	0.456	0.303	2.73	11	0	1	2	2
23	0.790	0.386	0.798	0.969	0.464	0.088	5.06	4	7	10	5	2
24	0.466	0.162	0.752	0.449	0.173	0.819	3.72	2	5	3	2	5
25	0.662	0.881	0.868	0.470	0.754	0.199	4.44	7	7	4	6	3
26	0.815	0.833	0.293	0.823	0.415	0.223	5.21	7	4	7	5	3
27	0.774	0.444	0.989	0.514	0.998	0.041	4.97	4	11	5	12	1
28	0.212	0.492	0.250	0.780	0.947	0.049	2.84	3	2	4	6	0
29	0.049	0.486	0.920	0.008	0.030	0.378	1.98	2	4	0	0	1
30	0.318	0.515	0.581	0.888	0.180	0.975	3.22	3	3	5	2	7
31	0.535	0.118	0.133	0.115	0.399	0.308	3.96	2	2	2	3	3
32	0.166	0.007	0.466	0.993	0.192	0.772	2.65	0	2	7	1	4
33	0.167	0.137	0.978	0.725	0.979	0.489	2.65	1	6	4	6	2
34	0.693	0.315	0.262	0.208	0.038	0.307	4.57	3	3	3	1	3
35	0.458	0.024	0.072	0.375	0.628	0.523	3.69	0	1	3	4	4
36	0.448	0.064	0.002	0.264	0.046	0.107	3.66	1	0	2	1	1
37	0.657	0.316	0.633	0.593	0.996	0.838	4.42	3	5	5	11	6
38	0.350	0.331	0.146	0.756	0.843	0.951	3.33	2	1	4	5	7
39	0.796	0.616	0.584	0.638	0.353	0.836	5.09	6	5	6	4	7
40	0.447	0.374	0.394	0.273	0.487	0.634	3.65	3	3	2	3	4
41	0.288	0.236	0.976	0.724	0.669	0.519	3.11	2	7	4	4	3
42	0.888	0.853	0.612	0.904	0.016	0.104	5.76	8	6	9	1	3
43	0.947	0.742	0.770	0.030	0.319	0.579	6.52	8	8	2	5	7
44	0.733	0.227	0.732	0.399	0.034	0.602	4.76	3	6	4	1	5
45	0.163	0.877	0.761	0.079	0.888	0.815	2.63	5	4	1	5	4
46	0.604	0.596	0.703	0.419	0.904	0.696	4.21	5	5	4	7	5
47	0.001	0.169	0.312	0.885	0.398	0.714	1.01	0	0	2	1	1
48	0.901	0.532	0.972	0.581	0.510	0.566	5.90	6	11	6	6	6
49	0.633	0.230	0.152	0.068	0.430	0.608	4.32	3	2	1	4	5
50	0.523	0.733	0.100	0.279	0.743	0.403	3.91	5	2	3	5	3

APPENDIX A2

COMPARING THE ABILITY OF THE CONVENTIONAL AND BAYESIAN METHODS TO ACCURATELY MEASURE THE TRUE LEVEL OF SAFETY.

Using the randomly generated accident frequencies from Appendix A1 the level of safety was estimated at each location for study periods ranging from 1 to 5 years using Conventional as well as Bayesian methods.

1. CONVENTIONAL ESTIMATES

The *conventional* estimates were determined as follows :

$$\bar{X}_1 = X_1$$

$$\bar{X}_2 = (X_1 + X_2)/2$$

$$\bar{X}_3 = (X_1 + X_2 + X_3)/3$$

$$\bar{X}_4 = (X_1 + X_2 + X_3 + X_4)/4$$

$$\bar{X}_5 = (X_1 + X_2 + X_3 + X_4 + X_5)/5$$

2. BAYESIAN ESTIMATES

$E(m)$ and $VAR(m)$ were determined from Year 1 data only, using the *method-of-sample-moments*: $E(m) = 4.05$ and $VAR(m) = 1.64$.

$$m_1 = a_1 E(m) + (1-a_1) X_1$$

$$a_1 = E(m)/[E(m) + VAR(m)] = 0.71$$

$$m_2 = a_2 E(m) + (1-a_2) \bar{X}_2$$

$$a_2 = E(m)/[E(m) + 2*VAR(m)] = 0.55$$

$$m_3 = a_3 E(m) + (1-a_3) \bar{X}_3$$

$$a_3 = E(m)/[E(m) + 3*VAR(m)] = 0.45$$

$$m_4 = a_4 E(m) + (1-a_4) \bar{X}_4$$

$$a_4 = E(m) / [E(m) + 4 \cdot \text{VAR}(m)] = 0.38$$

$$m_5 = a_5 E(m) + (1-a_5) \bar{X}_5$$

$$a_5 = E(m) / [E(m) + 5 \cdot \text{VAR}(m)] = 0.33$$

The Generic estimate of safety = E(m).

3. DEGREE OF DEVIATION

For each of these *Conventional* and *Bayesian* estimates the *absolute degree of deviation (D)* was determined as follows :

For Conventional estimates : $D_{i,y} = |m_{ti} - \bar{X}_{i,y}| * 100 / m_{ti}$

For Bayesian estimates : $D_{i,y} = |m_{ti} - m_{i,y}| * 100 / m_{ti}$

For the Generic estimate : $D_{i,y} = |m_{ti} - E(m)| * 100 / m_{ti}$

Where y = 1, 2,5 and i = 1, 2, 1000.

The *Conventional*, *Generic*, *Bayesian* estimates and the *degree-of-deviation* associated with each estimate at each site are indicated in Table A2.1

The next step was to organise the data into a cumulative frequency table as shown in Table A2.2. The Table shows, for each type and period of estimation, the % of sites smaller than a certain degree of deviation (D).

Table A2.1 : Conventional and Bayesian estimates and deviations from the 'true' level of safety. (Sites 1 to 50)

ID	m_t	X_1	D_1	X_2	D_2	X_3	D_3	X_4	D_4	X_5	D_5	D_g	m_1	D_{m1}	m_2	D_{m2}	m_3	D_{m3}	m_4	D_{m4}	m_5	D_{m5}
1	3.24	2	38.2	1.5	53.7	2.33	27.9	2.75	15.0	3.4	5.0	25.0	3.46	6.8	2.91	10.2	3.11	4.0	3.24	0.2	3.61	11.6
2	3.16	6	89.9	4.5	42.4	3.67	16.0	4	26.6	3.6	13.9	28.1	4.61	45.9	4.25	34.5	3.84	21.5	4.02	27.2	3.75	18.6
3	6.01	6	0.2	5.5	8.5	6	0.2	5.25	12.7	5.2	13.5	32.7	4.61	23.3	4.70	21.9	5.12	14.9	4.79	20.3	4.82	19.9
4	3.42	5	46.3	4.0	17.0	3.67	7.3	4	17.0	4.8	40.4	18.4	4.32	26.4	4.03	17.8	3.84	12.3	4.02	17.5	4.55	33.1
5	4.31	6	39.1	5.5	27.5	5.33	23.6	5.25	21.7	5.4	25.2	6.2	4.61	6.9	4.70	8.9	4.75	10.2	4.79	11.1	4.95	14.8
6	2.75	4	45.3	3.0	9.0	3.67	33.2	3.5	27.2	3.4	23.5	47.0	4.03	46.6	3.58	30.0	3.84	39.5	3.71	34.7	3.61	31.3
7	2.52	2	20.6	3.0	19.2	2.33	7.3	2	20.6	2.4	4.7	60.8	3.46	37.3	3.58	42.1	3.11	23.4	2.78	10.5	2.94	16.9
8	3.85	7	81.8	6.0	55.9	5	29.9	4.25	10.4	4.0	3.9	5.1	4.90	27.3	4.92	27.9	4.57	18.7	4.17	8.4	4.02	4.3
9	3.00	4	33.5	3.5	16.9	3.33	11.3	3	0.2	3.4	13.5	35.1	4.03	34.7	3.80	26.9	3.66	22.0	3.40	13.5	3.61	20.7
10	2.16	2	7.4	2.5	15.7	2.33	8.0	2.25	4.1	1.8	16.7	87.3	3.46	60.0	3.35	55.2	3.11	43.8	2.94	35.8	2.54	17.7
11	2.03	3	48.0	1.5	26.0	2	1.4	2.25	11.0	2.4	18.4	99.6	3.74	84.7	2.91	43.3	2.92	44.2	2.94	44.8	2.94	45.2
12	2.60	3	15.2	2.5	4.0	2	23.2	2.5	4.0	2.6	0.1	55.5	3.74	43.9	3.35	28.9	2.92	12.3	3.09	18.7	3.08	18.2
13	2.63	5	90.1	4.0	52.1	3.33	26.7	2.75	4.6	3.0	14.1	53.9	4.32	64.3	4.03	53.1	3.66	39.0	3.24	23.4	3.35	27.2
14	4.21	0	100	2.5	40.7	3	28.8	3	28.8	4.0	5.1	4.0	2.88	31.7	3.35	20.4	3.47	17.6	3.40	19.3	4.02	4.7
15	4.20	3	28.6	3.5	16.7	4.33	3.1	4.5	7.1	4.0	4.8	3.7	3.74	10.9	3.80	9.5	4.20	0.1	4.33	3.0	4.02	4.4
16	2.09	2	4.5	2.5	19.4	2.67	27.4	2.5	19.4	2.2	5.1	93.3	3.46	65.1	3.35	60.2	3.29	57.1	3.09	47.6	2.81	34.2
17	2.76	1	63.8	1.5	45.6	1.67	39.6	2	27.5	2.4	13.0	46.7	3.17	14.8	2.91	5.3	2.74	0.7	2.78	0.8	2.94	6.7
18	2.65	3	13.3	3.0	13.3	3	13.3	2.75	3.9	3.2	20.9	52.9	3.74	41.5	3.58	35.2	3.47	31.2	3.24	22.6	3.48	31.5
19	5.10	2	60.8	2.5	51.0	3	41.2	3	41.2	3.4	33.3	20.6	3.46	32.2	3.35	34.2	3.47	31.9	3.40	33.3	3.61	29.1
20	5.16	2	61.2	2.0	61.2	3.33	35.4	4.75	7.9	4.6	10.8	21.5	3.46	33.0	3.13	39.3	3.66	29.1	4.48	13.1	4.42	14.4
21	3.59	6	67.2	4.5	25.4	4.67	30.1	3.5	2.5	3.8	5.9	12.8	4.61	28.5	4.25	18.4	4.39	22.3	3.71	3.4	3.88	8.2
22	2.73	11	303	5.5	101	4	46.5	3.5	28.2	3.2	17.2	48.2	6.05	121.7	4.70	72.0	4.02	47.3	3.71	35.8	3.48	27.4
23	5.06	4	20.9	5.5	8.8	7	38.4	6.5	28.6	5.6	10.8	20.0	4.03	20.2	4.70	7.1	5.67	12.1	5.56	10.1	5.09	0.6
24	3.72	2	46.2	3.5	5.9	3.33	10.4	3	19.3	3.4	8.6	8.8	3.46	7.1	3.80	2.2	3.66	1.7	3.40	8.6	3.61	2.8
25	4.44	7	57.6	7.0	57.6	6	35.1	6	35.1	5.4	21.6	8.9	4.90	10.3	5.37	20.9	5.12	15.2	5.26	18.3	4.95	11.5

Continue

Table A2.1 (continued)

ID	m_t	X_1	D_1	X_2	D_2	X_3	D_3	X_4	D_4	X_5	D_5	D_g	m_1	D_{m1}	m_2	D_{m2}	m_3	D_{m3}	m_4	D_{m4}	m_5	D_{m5}
26	5.21	7	34.4	5.5	5.6	6	15.2	5.75	10.4	5.2	0.1	22.3	4.90	5.9	4.70	9.8	5.12	1.7	5.10	2.0	4.82	7.5
27	4.97	4	19.5	7.5	50.9	6.67	34.2	8	61.0	6.6	32.8	18.6	4.03	18.8	5.59	12.6	5.48	10.4	6.49	30.7	5.76	15.8
28	2.84	3	5.8	2.5	11.8	3	5.8	3.75	32.2	3.0	5.8	42.7	3.74	32.1	3.35	18.3	3.47	22.4	3.86	36.2	3.35	18.0
29	1.98	2	1.1	3.0	51.7	2	1.1	1.5	24.2	1.4	29.2	104.6	3.46	74.7	3.58	80.9	2.92	47.8	2.47	24.9	2.27	15.0
30	3.22	3	6.8	3.0	6.8	3.67	13.9	3.25	0.9	4.0	24.2	25.7	3.74	16.3	3.58	11.1	3.84	19.2	3.55	10.4	4.02	24.7
31	3.96	2	49.4	2.0	49.4	2	49.4	2.25	43.1	2.4	39.3	2.3	3.46	12.6	3.13	20.9	2.92	26.1	2.94	25.8	2.94	25.6
32	2.65	0	100	1.0	62.2	3	13.4	2.5	5.5	2.8	5.8	52.9	2.88	8.8	2.68	1.4	3.47	31.2	3.09	16.8	3.21	21.4
33	2.65	1	62.3	3.5	32.0	3.67	38.2	4.25	60.2	3.8	43.3	52.6	3.17	19.4	3.80	43.3	3.84	44.7	4.17	57.3	3.88	46.3
34	4.57	3	34.4	3.0	34.4	3	34.4	2.5	45.3	2.6	43.1	11.5	3.74	18.1	3.58	21.8	3.47	24.1	3.09	32.4	3.08	32.7
35	3.69	0	100.0	0.5	86.5	1.33	63.9	2	45.8	2.4	35.0	9.6	2.88	22.0	2.46	33.4	2.56	30.7	2.78	24.7	2.94	20.3
36	3.66	1	72.7	0.5	86.3	1	72.7	1	72.7	1.0	72.7	10.7	3.17	13.4	2.46	32.8	2.37	35.1	2.16	40.9	2.01	45.1
37	4.42	3	32.1	4.0	9.5	4.33	2.0	6	35.7	6.0	35.7	8.4	3.74	15.3	4.03	8.9	4.20	4.9	5.26	18.9	5.36	21.1
38	3.33	2	39.9	1.5	54.9	2.33	29.9	3	9.8	3.8	14.2	21.7	3.46	3.9	2.91	12.6	3.11	6.6	3.40	2.2	3.88	16.7
39	5.09	6	17.8	5.5	8.0	5.67	11.3	5.25	3.1	5.6	10.0	20.5	4.61	9.5	4.70	7.7	4.94	3.1	4.79	5.9	5.09	0.1
40	3.65	3	17.9	3.0	17.9	2.67	27.0	2.75	24.7	3.0	17.9	10.8	3.74	2.5	3.58	2.0	3.29	9.9	3.24	11.2	3.35	8.4
41	3.11	2	35.8	4.5	44.5	4.33	39.1	4.25	36.4	4.0	28.4	29.9	3.46	11.0	4.25	36.4	4.20	35.0	4.17	34.0	4.02	28.9
42	5.76	8	38.8	7.0	21.5	7.67	33.0	6	4.1	5.4	6.3	29.8	5.19	10.0	5.37	6.8	6.03	4.7	5.26	8.8	4.95	14.1
43	6.52	8	22.6	8.0	22.6	6	8.0	5.75	11.9	6.0	8.0	38.0	5.19	20.5	5.82	10.8	5.12	21.5	5.10	21.8	5.36	17.9
44	4.76	3	36.9	4.5	5.4	4.33	8.9	3.5	26.4	3.8	20.1	14.9	3.74	21.3	4.25	10.6	4.20	11.6	3.71	22.0	3.88	18.4
45	2.63	5	89.8	4.5	70.8	3.33	26.5	3.75	42.3	3.8	44.2	53.6	4.32	64.0	4.25	61.3	3.66	38.7	3.86	46.6	3.88	47.3
46	4.21	5	18.8	5.0	18.8	4.67	10.9	5.25	24.7	5.2	23.6	3.8	4.32	2.7	4.47	6.3	4.39	4.2	4.79	13.8	4.82	14.5
47	1.01	0	100.0	0.0	100.0	0.67	33.9	0.75	25.7	0.8	20.7	301.1	2.88	185.3	2.23	121.4	2.19	117.2	2.01	98.9	1.87	85.6
48	5.90	6	1.7	8.5	44.1	7.67	30.0	7.25	22.9	7.0	18.7	31.4	4.61	21.8	6.04	2.4	6.03	2.3	6.03	2.2	6.02	2.1
49	4.32	3	30.6	2.5	42.2	2	53.7	2.5	42.2	3.0	30.6	6.4	3.74	13.4	3.35	22.4	2.92	32.4	3.09	28.5	3.35	22.6
50	3.91	5	27.7	3.5	10.6	3.33	14.9	3.75	4.2	3.6	8.0	3.4	4.32	10.4	3.80	2.9	3.66	6.6	3.86	1.3	3.75	4.3

Table A2.2 : Cumulative frequencies (%) : Conventional, Generic and Bayesian estimates

D	X ₁	X ₂	X ₃	X ₄	X ₅	E(m)	m ₁	m ₂	m ₃	m ₄	m ₅
0	0	0	0	0	0	0	0	0	0	0	0
5	9.4	12.2	13.4	14.5	16.6	10.3	12.8	15.7	15.9	16.9	19.6
10	16.5	22.0	27.0	28.9	32.7	19.6	24.8	29.6	33.0	35.2	35.5
15	23.1	30.6	39.8	42.1	48.0	32.8	38.8	43.5	47.5	51.0	53.5
20	30.0	39.0	48.2	55.0	61.7	43.0	49.4	56.8	60.8	65.2	68.0
25	37.0	47.6	58.0	64.8	72.4	52.9	58.0	67.4	71.5	75.0	79.1
30	43.4	56.3	66.7	72.1	80.5	59.7	67.0	74.3	79.2	81.4	86.2
35	49.9	64.9	73.8	78.9	86.8	66.3	73.6	79.0	84.7	87.2	90.6
40	55.6	72.1	80.4	84.3	91.2	73.4	79.9	85.4	88.7	91.5	93.8
45	60.6	78.8	84.7	90.0	95.1	77.2	83.2	88.9	92.0	93.6	95.6
50	65.9	83.2	89.1	93.1	96.8	80.6	85.9	90.8	94.0	95.5	96.8
55	70.1	86.6	92.4	95.5	97.9	84.3	87.8	93.5	95.0	96.7	97.5
60	73.9	89.5	94.8	97.7	98.5	86.4	90.0	94.5	96.5	97.7	98.2
65	78.1	92.2	96.8	98.7	98.8	88.9	92.1	95.6	97.5	98.3	98.6
70	81.6	93.8	97.8	99.2	99.5	89.8	93.7	96.6	98.2	98.9	99.0
75	84.3	95.1	98.5	99.7	99.8	90.6	94.6	97.3	98.5	99.1	99.1
80	86.6	96.4	99.0	99.8	99.9	91.9	95.4	98.2	98.5	99.2	99.5
85	88.0	97.3	99.5	99.8	99.9	92.8	96.3	98.3	98.7	99.2	99.5
90	89.2	98.1	99.5	99.8	99.9	93.8	96.7	98.6	99.0	99.5	99.7
95	89.9	98.6	99.5	99.9	100.0	94.8	96.9	98.9	99.3	99.6	99.7
100	94.7	99.1	99.6	100.0	100.0	95.7	97.5	99.0	99.4	99.7	99.7
105	95.3	99.2	99.9	100.0	100.0	96.1	97.9	99.1	99.6	99.7	99.8
110	96.5	99.5	99.9	100.0	100.0	96.2	98.3	99.2	99.6	99.8	99.8
115	96.9	99.5	100.0	100.0	100.0	96.7	98.8	99.2	99.6	99.9	99.8
120	97.3	99.6	100.0	100.0	100.0	96.9	99.0	99.4	99.7	99.9	99.9
125	97.5	99.7	100.0	100.0	100.0	97.0	99.1	99.5	99.8	99.9	99.9
130	97.8	99.7	100.0	100.0	100.0	97.7	99.1	99.5	99.8	99.9	99.9
135	98.0	99.7	100.0	100.0	100.0	98.1	99.1	99.6	99.8	99.9	99.9
140	98.6	99.7	100.0	100.0	100.0	98.3	99.2	99.6	99.8	99.9	99.9
145	98.7	99.8	100.0	100.0	100.0	98.3	99.4	99.6	99.8	99.9	99.9
150	98.9	99.9	100.0	100.0	100.0	98.6	99.4	99.8	99.9	99.9	99.9
155	99.0	100.0	100.0	100.0	100.0	98.8	99.5	99.8	99.9	99.9	99.9
160	99.0	100.0	100.0	100.0	100.0	98.9	99.5	99.8	99.9	99.9	99.9
165	99.2	100.0	100.0	100.0	100.0	99.1	99.5	99.8	99.9	99.9	99.9
170	99.3	100.0	100.0	100.0	100.0	99.3	99.5	99.8	99.9	99.9	99.9
175	99.3	100.0	100.0	100.0	100.0	99.3	99.5	99.8	99.9	99.9	99.9
180	99.4	100.0	100.0	100.0	100.0	99.4	99.5	99.8	99.9	99.9	99.9
185	99.4	100.0	100.0	100.0	100.0	99.4	99.5	99.8	99.9	99.9	99.9
190	99.5	100.0	100.0	100.0	100.0	99.5	99.6	99.8	99.9	99.9	99.9
195	99.5	100.0	100.0	100.0	100.0	99.5	99.8	99.8	99.9	99.9	100.0
200	99.6	100.0	100.0	100.0	100.0	99.5	99.8	99.8	99.9	99.9	100.0

APPENDIX A3

EFFICIENCY ASSESSMENT OF IDENTIFICATION METHODS

In order to assess the efficiency of the CP1, CN2 and B1 identification methods the following procedures were applied :

- **Identifying 'true' hazardous locations.**

Previously the random number *rnd1* was generated for each of the 1000 sites. This random number was used to generate the true level of safety (m_t) from a Gamma distribution with mean = 4 acc/year and a variance = 2.

It is assumed that all those sites where $rnd1 > 0.95$ are 'true' hazardous locations. All the hazardous locations identified from the CP1, CN2 and B1 identification methods will be compared with these true hazardous locations to determine the efficiency of the respective methods.

Thus in Tables A3.1 and A3.2 : $H_t = 1$ of $rnd1 > 0.95$

- **Method CP1**

In Table A3.1 the value for P_{ij} (where $j = 1$ to 5) represent the probability that the accident total for a j year period exceeds the average accident total for the same period across all 1000 sites.

The average accident frequency for the 5 different study periods are shown below :

Study Period	1 Year	2 Year	3 Year	4 Year	5 Year
Average	4.05	8.04	12.02	16.08	20.2

If X_{ij} is the observed accident number at site i over a period of j years and Xa_j is the average accident number for a period of j years, then :

$$P_{ij} = \sum_{k=0}^{X_{ij}} \frac{(Xa_j)^k e^{-Xa_j}}{k!}$$

P_{ij} was determined using the following Microsoft® Excel® function:

$$P_{ij} = \text{POISSON}(X_{ij}, Xa_j, \text{TRUE})$$

A site was identified as hazardous if $P_j > 0.95$. Thus in Table A3.1 : $Hc_j = 1$ if $P_j > 0.95$.

The efficiency descriptions in the E_j columns were determined as follows :

'tp' – True positive : If $H_t = 1$ and $H_{pj} = 1$

'tn' – True negative : If $H_t = 0$ and $H_{pj} = 0$

'fp' – False positive : If $H_t = 0$ and $H_{pj} = 1$

'fn' – False negative : If $H_t = 1$ and $H_{pj} = 0$

- **Method CN2**

Table A3.1 contains the average accident frequency per location for each of the 5 study periods.

For each study period the critical accident number were estimated as follows :

$$Xcr_j = Xa_j + 1.645\sqrt{Xa_j} + 0.5$$

These critical values are as follows :

Study Period	1 Year	2 Year	3 Year	4 Year	5 Year
Critical Value	7.86	13.02	18.22	23.18	28.09

A site was identified as a hazardous location if $j^* \bar{X}_j > X_{crj}$, where X_j was obtained from Table A2.1 and $j = 1$ to 5.

Thus in Table A3.1 : $H_{jc} = 1$ if $j^* \bar{X}_j > X_{crj}$.

The efficiency descriptions in the E_{jc} columns were determined as follows :

'tp' – True positive : If $H_t = 1$ and $H_{jc} = 1$

'tn' – True negative : If $H_t = 0$ and $H_{jc} = 0$

'fp' – False positive : If $H_t = 0$ and $H_{jc} = 1$

'fn' – False negative : If $H_t = 1$ and $H_{jc} = 0$

• **Method B1**

In Table A3.2 m_{ij} refers to the Bayesian safety estimates that was determined for each site in Appendix A2.

The prior estimates $E(m)$ and $VAR(m)$ were determined from Year 1 data : $E(m) = 4.05$ and $VAR(m) = 1.64$. The Gamma parameters can be shown to be: $\alpha = 2.47$ and $\beta = 10$.

The posterior Gamma parameters of site i are determined as follows :

$$\alpha'_i = \alpha + j$$

$$\beta'_i = \beta + X_{Tij}$$

where

X_{Tij} - The total number of accidents at site i over a period of j years.

j - Length of study period : 1 to 5.

A site was identified as a hazardous location if $P(m_{ij} > E(m)) > 0.95$.

Where

$$P(m_{ij} > E(m)) = 1 - \int_0^{E(m)} \frac{\beta^{\alpha'} m^{\alpha'-1} e^{-\beta m_i}}{\Gamma(\alpha')} dm$$

The following Microsoft® Excel® function was used to calculate $P(m_{ij} > E(m))$:

$$= 1 - \text{GAMMADIST}[E(m), \beta', 1/\alpha', \text{TRUE}]$$

Thus in Table A3.2 : $P_{mj} = P(m_{ij} > E(m))$

$$H_{bj} = 1 \text{ if } P_{mj} > 0.95$$

The efficiency descriptions in the E_{bj} columns were determined as follows :

'tp' - True positive : If $H_t = 1$ and $H_{bj} = 1$

'tn' - True negative : If $H_t = 0$ and $H_{bj} = 0$

'fp' - False positive : If $H_t = 0$ and $H_{bj} = 1$

'fn' - False negative : If $H_t = 1$ and $H_{bj} = 0$

Table A3.1 : Methods CP1 and CN2 - Identification of hazardous locations (Sites 1 to 50)

ID	m	rnd1	H _m	P ₁	H _{p1}	E ₁	P ₂	H _{p2}	E ₂	P ₃	H _{p3}	E ₃	P ₄	H _{p4}	E ₄	P ₅	H _{p5}	E ₅	H _{1c}	E _{1c}	H _{1c}	E _{1c}	H _{1c}	E _{1c}	H _{1c}	E _{1c}	H _{1c}	E _{1c}
1	3.23	0.323	0	0.231	0	tn	0.041	0	tn	0.089	0	tn	0.123	0	tn	0.283	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
2	3.16	0.301	0	0.884	0	tn	0.712	0	tn	0.459	0	tn	0.558	0	tn	0.365	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
3	6.013	0.912	0	0.884	0	tn	0.885	0	tn	0.962	1	fp	0.907	0	tn	0.915	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
4	3.419	0.377	0	0.778	0	tn	0.588	0	tn	0.459	0	tn	0.558	0	tn	0.832	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
5	4.314	0.631	0	0.884	0	tn	0.885	0	tn	0.898	0	tn	0.907	0	tn	0.942	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
6	2.752	0.191	0	0.62	0	tn	0.309	0	tn	0.459	0	tn	0.36	0	tn	0.283	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
7	2.517	0.137	0	0.231	0	tn	0.309	0	tn	0.089	0	tn	0.021	0	tn	0.036	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
8	3.85	0.504	0	0.946	0	tn	0.934	0	tn	0.843	0	tn	0.652	0	tn	0.542	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
9	2.995	0.255	0	0.62	0	tn	0.448	0	tn	0.345	0	tn	0.188	0	tn	0.283	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
10	2.161	0.073	0	0.231	0	tn	0.188	0	tn	0.089	0	tn	0.042	0	tn	0.004	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
11	2.027	0.054	0	0.424	0	tn	0.041	0	tn	0.045	0	tn	0.042	0	tn	0.036	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
12	2.603	0.156	0	0.424	0	tn	0.188	0	tn	0.045	0	tn	0.075	0	tn	0.061	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
13	2.63	0.162	0	0.778	0	tn	0.588	0	tn	0.345	0	tn	0.123	0	tn	0.147	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
14	4.214	0.605	0	0.017	0	tn	0.188	0	tn	0.241	0	tn	0.188	0	tn	0.542	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
15	4.202	0.602	0	0.424	0	tn	0.448	0	tn	0.679	0	tn	0.736	0	tn	0.542	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
16	2.094	0.063	0	0.231	0	tn	0.188	0	tn	0.154	0	tn	0.075	0	tn	0.019	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
17	2.759	0.193	0	0.088	0	tn	0.041	0	tn	0.02	0	tn	0.021	0	tn	0.036	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
18	2.647	0.166	0	0.424	0	tn	0.309	0	tn	0.241	0	tn	0.123	0	tn	0.209	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
19	5.099	0.797	0	0.231	0	tn	0.188	0	tn	0.241	0	tn	0.188	0	tn	0.283	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
20	5.159	0.807	0	0.231	0	tn	0.098	0	tn	0.345	0	tn	0.807	0	tn	0.774	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
21	3.588	0.428	0	0.884	0	tn	0.712	0	tn	0.77	0	tn	0.36	0	tn	0.453	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
22	2.731	0.186	0	0.999	1	fp	0.885	0	tn	0.574	0	tn	0.36	0	tn	0.209	0	tn	1	fp	0	tn	0	tn	0	tn	0	tn
23	5.056	0.79	0	0.62	0	tn	0.885	0	tn	0.994	1	fp	0.992	1	fp	0.962	1	fp	0	tn	0	tn	1	fp	1	fp	0	tn
24	3.719	0.466	0	0.231	0	tn	0.448	0	tn	0.345	0	tn	0.188	0	tn	0.283	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn
25	4.443	0.662	0	0.946	0	tn	0.982	1	fp	0.962	1	fp	0.977	1	fp	0.942	0	tn	0	tn	1	fp	0	tn	1	fp	0	tn
26	5.207	0.815	0	0.946	0	tn	0.885	0	tn	0.962	1	fp	0.962	1	fp	0.915	0	tn	0	tn	0	tn	0	tn	0	tn	0	tn

Continue ...

Table A3.2 : Method B1 - Identification of hazardous locations (Sites 1 to 50).

ID	m _t	H _t	m ₁	P _{m1}	H _{b1}	E _{b1}	m ₂	P _{m2}	H _{b2}	E _{b2}	m ₃	P _{m3}	H _{m3}	E _{m3}	m ₄	P _{m4}	H _{m4}	E _{m4}	m ₅	P _{m5}	H _{m5}	E _{m5}
1	3.24	0	3.46	0.256	0	tn	2.91	0.088	0	tn	3.11	0.111	0	tn	3.24	0.131	0	tn	3.61	0.254	0	tn
2	3.16	0	4.61	0.665	0	tn	4.25	0.553	0	tn	3.84	0.376	0	tn	4.02	0.459	0	tn	3.75	0.318	0	tn
3	6.01	0	4.61	0.665	0	tn	4.70	0.723	0	tn	5.12	0.871	0	tn	4.79	0.803	0	tn	4.82	0.832	0	tn
4	3.42	0	4.32	0.566	0	tn	4.03	0.460	0	tn	3.84	0.376	0	tn	4.02	0.459	0	tn	4.55	0.730	0	tn
5	4.31	0	4.61	0.665	0	tn	4.70	0.723	0	tn	4.75	0.768	0	tn	4.79	0.803	0	tn	4.95	0.871	0	tn
6	2.75	0	4.03	0.459	0	tn	3.58	0.279	0	tn	3.84	0.376	0	tn	3.71	0.308	0	tn	3.61	0.254	0	tn
7	2.52	0	3.46	0.256	0	tn	3.58	0.279	0	tn	3.11	0.111	0	tn	2.78	0.038	0	tn	2.94	0.050	0	tn
8	3.85	0	4.90	0.752	0	tn	4.92	0.793	0	tn	4.57	0.701	0	tn	4.17	0.537	0	tn	4.02	0.459	0	tn
9	3.00	0	4.03	0.459	0	tn	3.80	0.367	0	tn	3.66	0.296	0	tn	3.40	0.181	0	tn	3.61	0.254	0	tn
10	2.16	0	3.46	0.256	0	tn	3.35	0.202	0	tn	3.11	0.111	0	tn	2.94	0.060	0	tn	2.54	0.012	0	tn
11	2.03	0	3.74	0.354	0	tn	2.91	0.088	0	tn	2.92	0.073	0	tn	2.94	0.060	0	tn	2.94	0.050	0	tn
12	2.60	0	3.74	0.354	0	tn	3.35	0.202	0	tn	2.92	0.073	0	tn	3.09	0.091	0	tn	3.08	0.075	0	tn
13	2.63	0	4.32	0.566	0	tn	4.03	0.460	0	tn	3.66	0.296	0	tn	3.24	0.131	0	tn	3.35	0.147	0	tn
14	4.21	0	2.88	0.107	0	tn	3.35	0.202	0	tn	3.47	0.224	0	tn	3.40	0.181	0	tn	4.02	0.459	0	tn
15	4.20	0	3.74	0.354	0	tn	3.80	0.367	0	tn	4.20	0.544	0	tn	4.33	0.613	0	tn	4.02	0.459	0	tn
16	2.09	0	3.46	0.256	0	tn	3.35	0.202	0	tn	3.29	0.162	0	tn	3.09	0.091	0	tn	2.81	0.032	0	tn
17	2.76	0	3.17	0.172	0	tn	2.91	0.088	0	tn	2.74	0.045	0	tn	2.78	0.038	0	tn	2.94	0.050	0	tn
18	2.65	0	3.74	0.354	0	tn	3.58	0.279	0	tn	3.47	0.224	0	tn	3.24	0.131	0	tn	3.48	0.196	0	tn
19	5.10	0	3.46	0.256	0	tn	3.35	0.202	0	tn	3.47	0.224	0	tn	3.40	0.181	0	tn	3.61	0.254	0	tn
20	5.16	0	3.46	0.256	0	tn	3.13	0.138	0	tn	3.66	0.296	0	tn	4.48	0.684	0	tn	4.42	0.669	0	tn
21	3.59	0	4.61	0.665	0	tn	4.25	0.553	0	tn	4.39	0.626	0	tn	3.71	0.308	0	tn	3.88	0.387	0	tn
22	2.73	0	6.05	0.951	1	fp	4.70	0.723	0	tn	4.02	0.460	0	tn	3.71	0.308	0	tn	3.48	0.196	0	tn
23	5.06	0	4.03	0.459	0	tn	4.70	0.723	0	tn	5.67	0.957	1	fp	5.56	0.961	1	fp	5.09	0.904	0	tn
24	3.72	0	3.46	0.256	0	tn	3.80	0.367	0	tn	3.66	0.296	0	tn	3.40	0.181	0	tn	3.61	0.254	0	tn
25	4.44	0	4.90	0.752	0	tn	5.37	0.895	0	tn	5.12	0.871	0	tn	5.26	0.919	0	tn	4.95	0.871	0	tn
26	5.21	0	4.90	0.752	0	tn	4.70	0.723	0	tn	5.12	0.871	0	tn	5.10	0.889	0	tn	4.82	0.832	0	tn
27	4.97	0	4.03	0.459	0	tn	5.59	0.929	0	tn	5.48	0.936	0	tn	6.49	0.997	1	fp	5.76	0.983	1	fp

Continue ...

Table A3.2 (continued)

ID	m_t	H_t	m_1	P_{m1}	H_{b1}	E_{b1}	m_2	P_{m2}	H_{b2}	E_{b2}	m_3	P_{m3}	H_{m3}	E_{m3}	m_4	P_{m4}	H_{m4}	E_{m4}	m_5	P_{m5}	H_{m5}	E_{m5}
28	2.84	0	3.74	0.354	0	tn	3.35	0.202	0	tn	3.47	0.224	0	tn	3.86	0.382	0	tn	3.35	0.147	0	tn
29	1.98	0	3.46	0.256	0	tn	3.58	0.279	0	tn	2.92	0.073	0	tn	2.47	0.013	0	tn	2.27	0.003	0	tn
30	3.22	0	3.74	0.354	0	tn	3.58	0.279	0	tn	3.84	0.376	0	tn	3.55	0.240	0	tn	4.02	0.459	0	tn
31	3.96	0	3.46	0.256	0	tn	3.13	0.138	0	tn	2.92	0.073	0	tn	2.94	0.060	0	tn	2.94	0.050	0	tn
32	2.65	0	2.88	0.107	0	tn	2.68	0.052	0	tn	3.47	0.224	0	tn	3.09	0.091	0	tn	3.21	0.107	0	tn
33	2.65	0	3.17	0.172	0	tn	3.80	0.367	0	tn	3.84	0.376	0	tn	4.17	0.537	0	tn	3.88	0.387	0	tn
34	4.57	0	3.74	0.354	0	tn	3.58	0.279	0	tn	3.47	0.224	0	tn	3.09	0.091	0	tn	3.08	0.075	0	tn
35	3.69	0	2.88	0.107	0	tn	2.46	0.029	0	tn	2.56	0.026	0	tn	2.78	0.038	0	tn	2.94	0.050	0	tn
36	3.66	0	3.17	0.172	0	tn	2.46	0.029	0	tn	2.37	0.014	0	tn	2.16	0.003	0	tn	2.01	0.001	0	tn
37	4.42	0	3.74	0.354	0	tn	4.03	0.460	0	tn	4.20	0.544	0	tn	5.26	0.919	0	tn	5.36	0.949	0	tn
38	3.33	0	3.46	0.256	0	tn	2.91	0.088	0	tn	3.11	0.111	0	tn	3.40	0.181	0	tn	3.88	0.387	0	tn
39	5.09	0	4.61	0.665	0	tn	4.70	0.723	0	tn	4.94	0.825	0	tn	4.79	0.803	0	tn	5.09	0.904	0	tn
40	3.65	0	3.74	0.354	0	tn	3.58	0.279	0	tn	3.29	0.162	0	tn	3.24	0.131	0	tn	3.35	0.147	0	tn
41	3.11	0	3.46	0.256	0	tn	4.25	0.553	0	tn	4.20	0.544	0	tn	4.17	0.537	0	tn	4.02	0.459	0	tn
42	5.76	0	5.19	0.824	0	tn	5.37	0.895	0	tn	6.03	0.982	1	fp	5.26	0.919	0	tn	4.95	0.871	0	tn
43	6.52	0	5.19	0.824	0	tn	5.82	0.953	1	fp	5.12	0.871	0	tn	5.10	0.889	0	tn	5.36	0.949	0	tn
44	4.76	0	3.74	0.354	0	tn	4.25	0.553	0	tn	4.20	0.544	0	tn	3.71	0.308	0	tn	3.88	0.387	0	tn
45	2.63	0	4.32	0.566	0	tn	4.25	0.553	0	tn	3.66	0.296	0	tn	3.86	0.382	0	tn	3.88	0.387	0	tn
46	4.21	0	4.32	0.566	0	tn	4.47	0.643	0	tn	4.39	0.626	0	tn	4.79	0.803	0	tn	4.82	0.832	0	tn
47	1.01	0	2.88	0.107	0	tn	2.23	0.015	0	tn	2.19	0.007	0	tn	2.01	0.002	0	tn	1.87	0.000	0	tn
48	5.90	0	4.61	0.665	0	tn	6.04	0.970	1	fp	6.03	0.982	1	fp	6.03	0.989	1	fp	6.02	0.993	1	fp
49	4.32	0	3.74	0.354	0	tn	3.35	0.202	0	tn	2.92	0.073	0	tn	3.09	0.091	0	tn	3.35	0.147	0	tn
50	3.91	0	4.32	0.566	0	tn	3.80	0.367	0	tn	3.66	0.296	0	tn	3.86	0.382	0	tn	3.75	0.318	0	tn

APPENDIX A4

ASSESSMENT OF EVALUATION METHODS

In Appendix A3 the CP1 and B1 methods were used to identify hazardous locations for study periods from 1 to 5 years.

For each site in a group of identified hazardous locations the true level of safety in the 'before' period (m_{tb}) was reduced by 20 % - this represents the 'true' effect of a safety treatment. Using the same procedure as in Appendix A1, 5 years of accident frequencies were randomly generated assuming a Poisson distribution around the true 'after' level of safety (m_{ta}).

Table A4.1 and Table A4.2 contains listings of the hazardous locations selected by Methods CP1 and B1 using 3 years of data, and information on the randomly generated accident frequencies for each of 5 years in the 'after' period.

Table A4.1 : Method CP1 – Hazardous locations : Period = 3 years

ID	T ₃	m _{bt}	m _{at}	Xa ₁	Xa ₂	Xa ₃	Xa ₄	Xa ₅
3	18	6.01	4.81	4	4	3	6	7
23	21	5.06	4.04	1	5	2	3	6
25	18	4.44	3.55	2	3	4	3	3
26	18	5.21	4.17	3	2	5	6	2
27	20	4.97	3.98	1	4	6	2	6
42	23	5.76	4.61	2	5	4	3	4
43	18	6.52	5.22	6	6	4	7	6
48	23	5.90	4.72	5	6	5	3	3
56	19	5.40	4.32	6	2	4	2	5
78	19	4.56	3.65	4	4	2	3	10
80	19	4.78	3.82	3	2	5	2	4
83	19	5.24	4.19	6	1	1	5	4
85	27	6.89	5.51	5	5	5	4	6
95	18	4.53	3.62	5	1	3	5	3
98	24	7.97	6.38	6	11	3	3	5
106	22	5.50	4.40	5	4	7	5	3
120	22	5.92	4.73	3	7	5	7	4
133	18	5.55	4.44	3	9	7	5	6
136	20	4.02	3.22	4	3	5	5	3
139	19	4.44	3.55	7	3	4	3	4
140	22	6.45	5.16	10	7	6	8	4
148	27	7.01	5.60	2	5	7	8	8
157	22	7.76	6.21	7	9	3	9	4
165	24	7.03	5.63	2	7	7	5	5
168	21	6.27	5.02	4	8	6	5	3
185	26	5.19	4.15	4	0	4	3	8
187	22	4.45	3.56	2	1	2	4	2
200	21	6.73	5.39	8	8	6	3	3
213	18	7.20	5.76	9	7	7	10	6
217	20	5.04	4.03	4	8	2	4	4
221	33	6.46	5.17	6	8	7	4	1
227	25	7.37	5.89	5	7	7	6	6
229	19	5.13	4.11	7	2	2	2	1
233	19	5.79	4.64	4	3	7	5	1
249	32	9.04	7.23	5	8	7	9	5
250	23	5.31	4.24	10	3	6	5	5
255	22	4.59	3.67	2	1	4	4	7
264	18	6.48	5.19	5	5	2	5	5
273	21	5.75	4.60	6	7	3	8	4
276	22	5.68	4.55	5	3	2	2	4
278	21	4.76	3.81	1	3	0	4	4
282	18	6.75	5.40	2	5	6	4	6
283	19	6.66	5.33	4	5	8	8	3
285	28	7.08	5.66	7	6	10	5	8
300	19	5.48	4.38	5	5	2	6	5
301	18	6.68	5.34	2	6	9	3	3
308	18	3.66	2.93	2	2	2	2	4

Table A4.1 (continued)

ID	T ₃	m _{bt}	m _{at}	Xa ₁	Xa ₂	Xa ₃	Xa ₄	Xa ₅
327	22	6.93	5.55	8	6	10	3	6
334	22	6.08	4.86	8	5	5	4	4
341	20	4.95	3.96	7	4	2	5	6
344	23	6.65	5.32	6	6	6	4	8
349	19	5.55	4.44	5	5	6	9	5
354	22	5.95	4.76	4	3	5	7	5
355	20	8.70	6.96	3	8	9	5	3
361	19	4.63	3.70	2	2	5	4	3
370	20	4.85	3.88	6	6	3	3	1
371	21	5.31	4.25	4	4	6	3	7
389	19	5.81	4.65	2	9	6	4	4
392	21	5.96	4.77	5	6	6	7	8
397	19	4.39	3.51	6	2	2	4	7
398	31	6.81	5.45	6	3	6	5	5
402	19	5.66	4.53	2	7	4	6	4
403	23	8.18	6.54	10	6	5	7	9
409	26	9.05	7.24	6	5	4	7	7
413	19	4.15	3.32	5	3	2	4	4
418	22	8.10	6.48	10	5	10	5	7
421	25	7.27	5.81	7	9	7	5	7
428	23	5.64	4.51	4	1	6	5	3
435	18	6.03	4.82	6	3	2	3	7
436	18	5.58	4.46	4	3	3	3	3
437	21	6.76	5.41	7	6	5	5	6
443	23	5.64	4.51	4	4	5	3	6
450	26	8.64	6.91	8	10	7	3	10
456	28	10.12	8.10	8	7	12	5	7
460	19	3.64	2.91	2	4	4	4	5
463	18	4.30	3.44	3	2	1	2	5
468	22	4.84	3.87	5	9	3	7	3
477	19	5.27	4.21	4	10	2	6	3
481	18	4.01	3.21	3	3	5	2	5
491	22	4.87	3.89	7	1	2	6	0
493	19	6.59	5.27	7	2	9	3	5
497	22	6.91	5.53	6	5	4	4	6
503	18	4.03	3.23	3	1	4	1	2
506	32	5.36	4.29	2	4	6	3	2
516	18	4.39	3.51	4	2	2	2	5
520	26	5.25	4.20	3	4	6	2	5
523	18	4.44	3.55	3	2	4	5	7
535	24	5.65	4.52	3	3	2	9	2
539	24	6.67	5.34	4	4	4	7	8
543	18	4.25	3.40	2	5	4	4	4
548	25	6.98	5.58	7	8	6	3	5
559	28	8.84	7.07	8	6	11	4	9
560	20	5.28	4.22	4	3	5	2	3
573	22	7.46	5.97	8	5	4	7	8
581	26	5.92	4.74	2	6	6	4	4

Table A4.1 (continued)

ID	T ₃	m _{bt}	m _{at}	Xa ₁	Xa ₂	Xa ₃	Xa ₄	Xa ₅
584	22	6.83	5.46	5	3	3	4	8
593	19	5.77	4.62	6	4	2	6	4
607	21	5.29	4.23	2	3	6	4	6
612	19	5.23	4.18	4	1	5	5	5
614	18	5.65	4.52	6	5	5	3	6
620	18	3.72	2.97	2	3	1	1	0
641	26	6.24	5.00	7	6	7	1	8
646	20	4.67	3.73	6	1	6	1	6
652	18	4.90	3.92	5	4	3	4	4
657	22	6.58	5.26	4	3	7	11	3
661	18	6.08	4.87	4	5	7	3	1
665	21	6.13	4.90	4	8	4	4	5
666	24	5.10	4.08	9	2	1	2	7
695	25	5.15	4.12	5	4	7	4	3
696	18	6.40	5.12	2	7	9	5	5
702	25	7.41	5.93	4	8	8	5	6
714	19	5.79	4.63	3	4	4	4	4
715	23	5.31	4.25	3	5	6	4	5
719	22	5.07	4.06	4	5	3	4	3
730	27	7.42	5.94	0	3	3	6	4
750	24	5.72	4.58	3	8	5	4	3
767	24	4.97	3.98	2	1	2	2	8
770	18	3.60	2.88	1	3	2	2	3
780	21	7.33	5.86	6	10	7	6	8
782	18	5.87	4.70	1	2	4	5	6
786	19	4.82	3.86	4	6	5	3	4
789	18	6.94	5.55	7	2	4	6	2
808	18	4.09	3.27	3	3	2	10	8
816	18	6.24	4.99	7	2	6	10	3
822	22	5.82	4.65	2	5	7	2	8
827	18	6.65	5.32	7	3	5	2	9
851	21	5.32	4.25	5	2	2	4	2
856	19	4.19	3.35	6	2	2	3	4
859	24	6.56	5.25	4	9	4	4	2
873	18	3.87	3.10	2	2	4	4	4
881	32	6.45	5.16	4	5	2	3	5
885	18	4.37	3.50	5	5	2	4	9
889	19	4.84	3.87	9	4	3	1	5
892	18	5.76	4.61	8	5	12	5	6
893	18	4.91	3.93	9	5	5	3	5
895	22	4.08	3.26	2	2	0	7	6
899	18	7.89	6.31	10	10	8	9	10
902	23	4.58	3.67	6	5	2	5	3
903	23	6.16	4.93	7	3	3	6	7
904	18	5.78	4.62	4	9	5	4	4
913	20	4.72	3.78	2	6	4	2	4
914	21	5.90	4.72	4	5	3	7	4
926	18	5.23	4.18	2	7	6	8	9

Table A4.1 (continue)

ID	T ₃	m _{bt}	m _{at}	Xa ₁	Xa ₂	Xa ₃	Xa ₄	Xa ₅
932	22	6.08	4.86	5	7	5	7	6
934	23	8.45	6.76	7	8	5	3	6
935	18	6.22	4.98	7	6	8	8	3
943	26	7.42	5.94	4	3	6	4	6
950	18	6.73	5.38	9	7	4	4	7
951	24	6.42	5.13	6	5	3	6	5
955	20	6.99	5.59	7	6	9	7	7
961	22	4.97	3.97	5	2	7	1	9
962	28	6.57	5.25	6	6	7	8	5
963	18	5.06	4.05	2	1	6	4	2
964	18	5.79	4.63	2	3	4	8	4
965	19	4.53	3.63	3	5	4	0	6
966	23	6.24	4.99	7	3	4	3	2
972	27	4.88	3.90	6	5	6	9	4
991	19	4.72	3.77	2	2	2	5	6
994	18	5.36	4.28	4	6	5	1	3
Total	3386	2779.81	2223.84	753	743	759	728	786

Table A4.2 : Method B1 – hazardous locations : Period = 3 years

ID	m ₃	m _{bt}	m _{at}	Xa ₁	Xa ₂	Xa ₃	Xa ₄	Xa ₅
23	5.67	5.06	4.04	1	5	2	3	6
42	6.03	5.76	4.61	2	5	4	3	4
48	6.03	5.90	4.72	5	6	5	3	3
85	6.77	6.89	5.51	5	5	5	4	6
98	6.22	7.97	6.38	6	11	3	3	5
106	5.85	5.50	4.40	5	4	7	5	3
120	5.85	5.92	4.73	3	7	5	7	4
140	5.85	6.45	5.16	10	7	6	8	4
148	6.77	7.01	5.60	2	5	7	8	8
157	5.85	7.76	6.21	7	9	3	9	4
165	6.22	7.03	5.63	2	7	7	5	5
168	5.67	6.27	5.02	4	8	6	5	3
185	6.58	5.19	4.15	4	0	4	3	8
187	5.85	4.45	3.56	2	1	2	4	2
200	5.67	6.73	5.39	8	8	6	3	3
221	7.86	6.46	5.17	6	8	7	4	1
227	6.40	7.37	5.89	5	7	7	6	6
249	7.68	9.04	7.23	5	8	7	9	5
250	6.03	5.31	4.24	10	3	6	5	5
255	5.85	4.59	3.67	2	1	4	4	7
273	5.67	5.75	4.60	6	7	3	8	4
276	5.85	5.68	4.55	5	3	2	2	4
278	5.67	4.76	3.81	1	3	0	4	4
285	6.95	7.08	5.66	7	6	10	5	8
327	5.85	6.93	5.55	8	6	10	3	6
334	5.85	6.08	4.86	8	5	5	4	4
344	6.03	6.65	5.32	6	6	6	4	8
354	5.85	5.95	4.76	4	3	5	7	5
371	5.67	5.31	4.25	4	4	6	3	7

Table A4.2 (continued)

ID	m_3	m_{bt}	m_{at}	Xa_1	Xa_2	Xa_3	Xa_4	Xa_5
392	5.67	5.96	4.77	5	6	6	7	8
398	7.50	6.81	5.45	6	3	6	5	5
403	6.03	8.18	6.54	10	6	5	7	9
409	6.58	9.05	7.24	6	5	4	7	7
418	5.85	8.10	6.48	10	5	10	5	7
421	6.40	7.27	5.81	7	9	7	5	7
428	6.03	5.64	4.51	4	1	6	5	3
437	5.67	6.76	5.41	7	6	5	5	6
443	6.03	5.64	4.51	4	4	5	3	6
450	6.58	8.64	6.91	8	10	7	3	10
456	6.95	10.12	8.10	8	7	12	5	7
468	5.85	4.84	3.87	5	9	3	7	3
491	5.85	4.87	3.89	7	1	2	6	0
497	5.85	6.91	5.53	6	5	4	4	6
506	7.68	5.36	4.29	2	4	6	3	2
520	6.58	5.25	4.20	3	4	6	2	5
535	6.22	5.65	4.52	3	3	2	9	2
539	6.22	6.67	5.34	4	4	4	7	8
548	6.40	6.98	5.58	7	8	6	3	5
559	6.95	8.84	7.07	8	6	11	4	9
573	5.85	7.46	5.97	8	5	4	7	8
581	6.58	5.92	4.74	2	6	6	4	4
584	5.85	6.83	5.46	5	3	3	4	8
607	5.67	5.29	4.23	2	3	6	4	6
641	6.58	6.24	5.00	7	6	7	1	8
657	5.85	6.58	5.26	4	3	7	11	3
665	5.67	6.13	4.90	4	8	4	4	5
666	6.22	5.10	4.08	9	2	1	2	7
695	6.40	5.15	4.12	5	4	7	4	3
702	6.40	7.41	5.93	4	8	8	5	6
715	6.03	5.31	4.25	3	5	6	4	5
719	5.85	5.07	4.06	4	5	3	4	3
730	6.77	7.42	5.94	0	3	3	6	4
750	6.22	5.72	4.58	3	8	5	4	3
767	6.22	4.97	3.98	2	1	2	2	8
780	5.67	7.33	5.86	6	10	7	6	8
822	5.85	5.82	4.65	2	5	7	2	8
851	5.67	5.32	4.25	5	2	2	4	2
859	6.22	6.56	5.25	4	9	4	4	2
881	7.68	6.45	5.16	4	5	2	3	5
895	5.85	4.08	3.26	2	2	0	7	6
902	6.03	4.58	3.67	6	5	2	5	3
903	6.03	6.16	4.93	7	3	3	6	7
914	5.67	5.90	4.72	4	5	3	7	4
932	5.85	6.08	4.86	5	7	5	7	6
934	6.03	8.45	6.76	7	8	5	3	6
943	6.58	7.42	5.94	4	3	6	4	6
951	6.22	6.42	5.13	6	5	3	6	5

Table A4.2 (continued)

ID	m₃	m_{bt}	m_{at}	Xa₁	Xa₂	Xa₃	Xa₄	Xa₅
961	5.85	4.97	3.97	5	2	7	1	9
962	6.95	6.57	5.25	6	6	7	8	5
966	6.03	6.24	4.99	7	3	4	3	2
972	6.77	4.88	3.90	6	5	6	9	4
Total	502.10	512.18	409.75	411	419	412	394	426

APPENDIX B1

THE METHOD OF SAMPLE MOMENTS

In Table B1.1 is listed all the road links in the Province of the Western Cape, South Africa, that meet the following criteria. :

- Number of lanes : 2 lanes (one in each direction)
- Type of shoulder : Surfaced
- Surfaced width : ≥ 11 m (typically 2*3.7 m lanes with shoulder widths > 1.8 m)
- Geography : Rural – Province of the Western Cape
- AADT : > 1000 veh/day and < 8000 veh/day
- Segment length : > 500 m
- Accident period : 1993 to 1996 (4 Years)
- Accident type : ALL

This information was obtained from the road network and accident databases of the Transport Branch of the Department of Economic Affairs, Agriculture and Tourism of the Provincial Administration of the Western Cape

The *method of sample moments* has been applied as follows to estimate the true level of safety at each of these segments :

Mean accident rate (\bar{R}) = 1.15

Variance of \bar{R} (σ^2) = 1.13

Harmonic mean (E^*) = 9.60

$E(m) = 1.15$ and $VAR(m) = [9.60(1.13)-1.15]/9.60 = 1.01$

$\alpha = 9.60(1.15)/[1.13(9.60) - 1.15] = 1.14$ and $\beta = 1.14*1.15 = 1.31$

$\alpha'_i = 1.14 + E_i$ and $\beta' = 1.31 + A_i$

$m_i = \beta'_i/\alpha'_i$ and $VAR(m_i) = \beta'_i/(\alpha'_i)^2$

Table B1.1 : Method of sample moments

Road	Start	End	L	AADT	E	A	R	α'	β'	m	VAR(m)
NR00205	51.88	52.6	0.74	2704	2.92	23	7.87	4.06	24.31	5.98	1.47
MR00165	0	3.63	3.63	1502	7.97	32	4.02	9.11	33.31	3.66	0.40
MR00027	51.73	52.3	0.56	4220	3.45	15	4.34	4.59	16.31	3.55	0.77
MR00227	5.89	9.66	3.77	1111	6.12	21	3.43	7.26	22.31	3.07	0.42
TR02801	0	2.14	2.14	1158	3.62	12	3.31	4.76	13.31	2.80	0.59
TR03201	44.35	45.2	0.8	1760	2.06	7	3.40	3.20	8.31	2.60	0.81
MR00027	51.15	51.7	0.58	7077	6.00	17	2.83	7.14	18.31	2.57	0.36
MR00027	67.19	68.6	1.37	4722	9.45	25	2.65	10.59	26.31	2.48	0.23
MR00223	6.3	9.63	3.33	1618	7.87	21	2.67	9.01	22.31	2.48	0.27
MR00165	3.63	7.47	3.84	5530	31.02	76	2.45	32.17	77.31	2.40	0.07
TR03201	42.84	44.4	1.51	1708	3.77	10	2.65	4.91	11.31	2.31	0.47
NR00108	2.68	3.61	0.93	2530	3.44	9	2.62	4.58	10.31	2.25	0.49
NR00205	9.85	11.4	1.55	5593	12.67	29	2.29	13.81	30.31	2.20	0.16
TR03201	0	5.71	5.71	1729	14.42	32	2.22	15.56	33.31	2.14	0.14
NR00205	49.34	51.8	2.46	4107	14.76	32	2.17	15.90	33.31	2.10	0.13
TR02801	2.14	3.73	1.59	3762	8.74	19	2.17	9.88	20.31	2.06	0.21
TR00204	50.54	55	4.49	2663	17.47	34	1.95	18.61	35.31	1.90	0.10
NR00205	52.62	58.6	5.98	3447	30.12	47	1.56	31.26	48.31	1.55	0.05
NR00205	40.64	49.3	8.7	3063	38.93	58	1.49	40.07	59.31	1.48	0.04
TR00202	37.09	42.4	5.27	7300	56.21	83	1.48	57.35	84.31	1.47	0.03
DR01056	0	1.32	1.32	1108	2.14	5	2.34	3.28	6.31	1.93	0.59
MR00313	1.6	3.1	1.5	1843	4.04	8	1.98	5.18	9.31	1.80	0.35
TR00204	44.28	45.2	0.88	2449	3.15	6	1.91	4.29	7.31	1.71	0.40
TR02801	16.37	17.2	0.81	3850	4.56	8	1.76	5.70	9.31	1.64	0.29
TR02801	17.18	19.9	2.75	4029	16.19	25	1.54	17.33	26.31	1.52	0.09
NR00205	58.6	59.2	0.61	3519	3.14	5	1.59	4.28	6.31	1.48	0.35
TR03302	2.74	5.01	2.27	1365	4.53	7	1.55	5.67	8.31	1.47	0.26
TR02801	19.93	23.6	3.67	4240	22.73	33	1.45	23.87	34.31	1.44	0.06
NR00205	0	2.02	2.02	4214	12.44	18	1.45	13.58	19.31	1.42	0.10
DR01101	1.9	5.79	3.89	1980	11.25	15	1.33	12.39	16.31	1.32	0.11
TR00204	2.84	4.4	1.56	2981	6.79	9	1.32	7.93	10.31	1.30	0.16
NR00105	29.83	31.9	2.06	2610	7.86	10	1.27	9.00	11.31	1.26	0.14
DR01105	0	3.99	3.99	6717	39.16	49	1.25	40.30	50.31	1.25	0.03
TR02801	3.73	9.1	5.37	3031	23.78	29	1.22	24.92	30.31	1.22	0.05
MR00216	0	3.45	3.45	4520	22.78	27	1.19	23.92	28.31	1.18	0.05
TR00204	45.16	50.5	5.38	2492	19.59	23	1.17	20.73	24.31	1.17	0.06
TR02801	23.6	25.8	2.23	4803	15.65	18	1.15	16.79	19.31	1.15	0.07
DR01101	0	1.9	1.9	3779	10.49	12	1.14	11.63	13.31	1.14	0.10
TR02901	39.92	45.2	5.29	1045	8.08	9	1.11	9.22	10.31	1.12	0.12
MR00279	33.98	34.7	0.67	2016	1.97	2	1.01	3.11	3.31	1.06	0.34
TR02901	51.85	55.1	3.23	1060	5.00	5	1.00	6.14	6.31	1.03	0.17
MR00279	34.65	35.3	0.68	3142	3.12	3	0.96	4.26	4.31	1.01	0.24
NR00205	7.12	9.12	2	6182	18.06	18	1.00	19.20	19.31	1.01	0.05
MR00187	6.96	8.03	1.07	5894	9.21	9	0.98	10.35	10.31	1.00	0.10
TR03302	1.74	2.74	1	1507	2.20	2	0.91	3.34	3.31	0.99	0.30
NR00208	65.71	67.8	2.07	5417	16.38	15	0.92	17.52	16.31	0.93	0.05
TR03201	22.79	26.2	3.42	1564	7.81	7	0.90	8.96	8.31	0.93	0.10
NR00107	0	34.9	34.86	2526	128.65	117	0.91	129.79	118.31	0.91	0.01
TR03201	5.71	15.3	9.55	1566	21.85	19	0.87	22.99	20.31	0.88	0.04

Continue ...

Table B1.1 (continued)

Road	Start	End	L	AADT	E	A	R	α'	β'	m	VAR(m)
TR03201	15.26	22.8	7.53	1372	15.09	13	0.86	16.23	14.31	0.88	0.05
MR00191	16.66	20.2	3.54	3808	19.69	17	0.86	20.84	18.31	0.88	0.04
NR00108	62	72.4	10.44	2465	37.60	32	0.85	38.74	33.31	0.86	0.02
MR00227	1	5.89	4.89	1359	9.71	8	0.82	10.85	9.31	0.86	0.08
NR00205	9.12	9.85	0.73	5839	6.23	5	0.80	7.37	6.31	0.86	0.12
MR00166	3.93	4.71	0.78	1405	1.60	1	0.62	2.74	2.31	0.84	0.31
TR03102	1.94	9.04	7.1	3656	37.92	31	0.82	39.06	32.31	0.83	0.02
TR00202	54.11	58.2	4.1	3524	21.11	17	0.81	22.25	18.31	0.82	0.04
NR00206	35.76	37.2	1.46	3647	7.78	6	0.77	8.92	7.31	0.82	0.09
NR00108	56.93	62	5.07	2550	18.89	15	0.79	20.03	16.31	0.81	0.04
MR00027	42.91	44.8	1.9	7270	20.18	16	0.79	21.32	17.31	0.81	0.04
MR00187	4.36	6.96	2.6	6407	24.34	19	0.78	25.48	20.31	0.80	0.03
MR00027	44.81	50.7	5.86	6922	59.26	45	0.76	60.40	46.31	0.77	0.01
NR00108	39.85	56.9	17.08	2277	56.82	43	0.76	57.96	44.31	0.76	0.01
TR03102	10.78	13.6	2.8	2722	11.14	8	0.72	12.28	9.31	0.76	0.06
TR03103	1.9	18.4	16.53	1284	31.01	23	0.74	32.15	24.31	0.76	0.02
MR00187	8.03	14.7	6.68	6170	60.22	45	0.75	61.36	46.31	0.75	0.01
NR00205	32.72	40.6	7.92	3386	39.18	29	0.74	40.32	30.31	0.75	0.02
TR02801	9.1	16.4	7.27	2831	30.07	22	0.73	31.21	23.31	0.75	0.02
TR03201	32.73	42.8	10.11	1563	23.09	16	0.69	24.23	17.31	0.71	0.03
TR00202	52.64	54.1	1.47	4247	9.12	6	0.66	10.26	7.31	0.71	0.07
MR00027	68.56	72.1	3.57	6622	34.54	23	0.67	35.68	24.31	0.68	0.02
MR00188	16.54	19.2	2.63	2540	9.76	6	0.61	10.90	7.31	0.67	0.06
NR00106	0.42	29.9	29.46	2514	108.21	72	0.67	109.35	73.31	0.67	0.01
TR02901	24.03	27.3	3.3	1129	5.44	3	0.55	6.58	4.31	0.66	0.10
NR00208	0	6.35	6.35	7379	68.46	44	0.64	69.60	45.31	0.65	0.01
MR00344	0	6.2	6.2	2247	20.35	12	0.59	21.49	13.31	0.62	0.03
MR00191	20.2	20.7	0.5	3576	2.61	1	0.38	3.75	2.31	0.62	0.16
NR00208	46.62	48.6	2.01	5398	15.85	9	0.57	16.99	10.31	0.61	0.04
NR00107	61.64	68.3	6.64	2839	27.54	16	0.58	28.68	17.31	0.60	0.02
NR00206	25.35	27.7	2.32	3770	12.78	7	0.55	13.92	8.31	0.60	0.04
NR00105	81.64	84.5	2.82	2701	11.13	6	0.54	12.27	7.31	0.60	0.05
NR00205	2.02	7.12	5.1	6241	46.50	27	0.58	47.64	28.31	0.59	0.01
TR00203	26.1	38.2	12.14	2064	36.61	21	0.57	37.75	22.31	0.59	0.02
NR00206	16.61	25.4	8.74	2734	34.91	19	0.54	36.05	20.31	0.56	0.02
NR00205	59.21	66.8	7.54	3706	40.83	22	0.54	41.97	23.31	0.56	0.01
NR00108	3.61	39.9	36.24	2602	137.77	75	0.54	138.91	76.31	0.55	0.00
NR00206	58.17	62.4	4.26	3721	23.16	12	0.52	24.30	13.31	0.55	0.02
NR00205	70.77	82.1	11.28	4187	69.00	37	0.54	70.14	38.31	0.55	0.01
TR00202	58.21	64.3	6.07	4694	41.63	22	0.53	42.77	23.31	0.55	0.01
NR00107	34.86	61.6	26.78	2665	104.27	56	0.54	105.41	57.31	0.54	0.01
NR00206	46.84	58.2	11.33	4114	68.10	36	0.53	69.24	37.31	0.54	0.01
DR01105	3.99	5.12	1.13	4461	7.36	3	0.41	8.51	4.31	0.51	0.06
NR00206	38.65	45.1	6.43	3327	31.25	15	0.48	32.40	16.31	0.50	0.02
TR07501	29.55	31.5	1.94	2640	7.48	3	0.40	8.62	4.31	0.50	0.06
NR00208	63.69	65.7	2.02	6003	17.72	8	0.45	18.86	9.31	0.49	0.03
NR00206	45.08	46.8	1.76	3189	8.20	3	0.37	9.34	4.31	0.46	0.05
NR00105	31.89	39.8	7.88	2715	31.26	13	0.42	32.40	14.31	0.44	0.01

Continue ...

Table B1.1 (continued)

Road	Start	End	L	AADT	E	A	R	α'	β'	m	VAR(m)
NR00104	27.78	54.2	26.46	3609	139.52	60	0.43	140.66	61.31	0.44	0.00
MR00177	18.12	22.5	4.39	7495	48.07	20	0.42	49.21	21.31	0.43	0.01
NR00105	39.77	81.6	41.87	2617	160.09	68	0.42	161.23	69.31	0.43	0.00
MR00200	3.64	11.3	7.7	7465	83.98	34	0.40	85.12	35.31	0.41	0.00
MR00223	0	6.3	6.3	4015	36.96	14	0.38	38.10	15.31	0.40	0.01
NR00206	12.44	16.6	4.17	2802	17.07	6	0.35	18.21	7.31	0.40	0.02
NR00206	62.43	68.3	5.87	3193	27.38	10	0.37	28.52	11.31	0.40	0.01
NR00205	66.75	70.8	4.02	3432	20.16	7	0.35	21.30	8.31	0.39	0.02
NR00208	48.63	54.9	6.28	6260	57.44	18	0.31	58.58	19.31	0.33	0.01
NR00208	8.27	16.8	8.51	5961	74.11	23	0.31	75.25	24.31	0.32	0.00
NR00206	68.3	72.9	4.55	4092	27.20	7	0.26	28.34	8.31	0.29	0.01
NR00205	11.4	14.7	3.27	5096	24.35	6	0.25	25.49	7.31	0.29	0.01
MR00199	19.57	22.5	2.89	3751	15.84	3	0.19	16.98	4.31	0.25	0.01
NR00207	43.29	52.1	8.78	7415	95.12	23	0.24	96.26	24.31	0.25	0.00
NR00208	59.42	63.7	4.27	6011	37.50	8	0.21	38.64	9.31	0.24	0.01
NR00208	34.32	46.6	12.3	6074	109.15	23	0.21	110.29	24.31	0.22	0.00

LEGEND

Road : The number assigned to the road by the provincial roads authority where NR = National Road, TR = Trunk Road and MR = Main Road.

Start : The kilometre value (according to the provincial road logs) at which the segment start.

End : The kilometre value (according to the provincial road logs) at which the segment end.

L : The length of the segment (End – Start).

AADT : Annual Average Daily Traffic

E : Exposure in million vehicle kilometres (= $L*365*4*AADT*10^{-6}$)

A : The total number of accidents recorded on the segment between 1993 and 1996 (4 years)

R : The accident rate (accidents per million vehicle kilometres) : = A/E .

α' : The posterior alpha parameter of the Gamma distribution of $m_i = \alpha + E$

β' : The posterior beta parameter of the Gamma distribution of $m_i = \beta + A$

m_i : The posterior Bayesian estimate = β'/α'

var(m_i) : The variance of the posterior Bayesian estimate = $\beta'/(\alpha')^2$

APPENDIX B2

IDENTIFICATION OF HAZARDOUS LOCATIONS

The 113 segments identified in Appendix B1 were evaluated using the following hazardous location identification methods :

- **The accident number method (CN1)**

The sample mean accident number (\bar{X}) = 5.57 acc/km

The sample standard deviation (σ_x) = 5.41

From Eqn. 5.1 $X_{cr} = 5.57 + 1.645 \cdot 5.41 = 14.47$

In Table B2.1 the value of CN1 = 1 if $N > 14.47$.

- **The accident rate method (CR1)**

The sample mean accident rate (\bar{R}) = 1.15 acc/mvkm

The sample standard deviation (σ_R) = 1.06

From Eqn. 5.4 $R_{cr} = 1.15 + 1.645 \cdot 1.06 = 2.90$

In Table B2.1 the value of CR1 = 1 if $R > 2.90$

- **The rate-number method (RN)**

In Table B2.1 the value of RN = 1 if both CN1 and CR1 = 1

- **The rate quality control method (CR2)**

The total number of accidents recorded on 113 links = 2433.

The sum of all exposures on 113 links = 3270.29 mvkm.

From Eqn 5.9 $R_a = 2433/3270.29 = 0.74$ acc/mvkm

In Table B2.1 $R_c = 0.74 + 1.645(0.74/E) + 0.5/E$ and $CR2 = 1$ if $R > R_c$.

- **The B1 Bayesian method**

In Table B2.1 $P_{b1} = P(m > 1.15)$ which was determined using the following Microsoft® Excel® function :

=1-GAMMADIST(1.15, β ,1/ α ,TRUE) where α and β are the Gamma parameters determined in Appendix B1.

In Table B2.1 $B1 = 1$ if $P_{b1} > 0.95$

- **The B2 Bayesian method**

In Table B2.1 $P_{21} = P(m > 0.74)$ which was determined using the following Microsoft® Excel® function :

=1-GAMMADIST(0.74, β ,1/ α ,TRUE) where α and β are the Gamma parameters determined in Appendix B1.

In Table B2.1 $B2 = 1$ if $P_{b2} > 0.95$

Table B2.1 : Identification of hazardous locations

Road	Start	End	E	N	R	m	CN1	CR1	NR	R _c	CR2	P _{b1}	B1	P _{b2}	B2
NR00205	51.88	52.62	2.92	31.08	7.87	6.15	1	1	1	1.80	1	1.000	1	1.000	1
MR00027	51.73	52.29	3.45	26.79	4.34	3.77	1	1	1	1.69	1	1.000	1	1.000	1
MR00165	0	3.63	7.97	8.82	4.02	2.92	0	1	0	1.32	1	0.994	1	0.997	1
MR00227	5.89	9.66	6.12	5.57	3.43	2.39	0	1	0	1.41	1	0.946	0	0.965	1
TR03201	44.35	45.15	2.06	8.75	3.40	2.62	0	1	0	2.12	1	0.987	1	0.993	1
TR02801	0	2.14	3.62	5.61	3.31	2.35	0	1	0	1.66	1	0.942	0	0.963	1
MR00027	51.15	51.73	6.00	29.31	2.83	2.64	1	0	0	1.41	1	1.000	1	1.000	1
MR00223	6.3	9.63	7.87	6.31	2.67	2.11	0	0	0	1.33	1	0.922	0	0.950	1
TR03201	42.84	44.35	3.77	6.62	2.65	2.12	0	0	0	1.63	1	0.929	0	0.955	1
MR00027	67.19	68.56	9.45	18.25	2.65	2.40	1	0	0	1.28	1	0.998	1	0.999	1
NR00108	2.68	3.61	3.44	9.68	2.62	2.22	0	0	0	1.69	1	0.970	1	0.984	1
MR00165	3.63	7.47	31.02	19.79	2.45	2.26	1	0	0	1.12	1	0.997	1	0.999	1
DR01056	0	1.32	2.14	3.79	2.34	1.80	0	0	0	2.08	1	0.790	0	0.842	0
NR00205	9.85	11.4	12.67	18.71	2.29	2.13	1	0	0	1.22	1	0.992	1	0.997	1
TR03201	0	5.71	14.42	5.60	2.22	1.85	0	0	0	1.20	1	0.851	0	0.897	0
TR02801	2.14	3.73	8.74	11.95	2.17	1.97	0	0	0	1.30	1	0.957	1	0.978	1
NR00205	49.34	51.8	14.76	13.01	2.17	1.98	0	0	0	1.19	1	0.964	1	0.983	1
MR00313	1.6	3.1	4.04	5.33	1.98	1.70	0	0	0	1.59	1	0.797	0	0.854	0
TR00204	50.54	55.03	17.47	7.57	1.95	1.74	0	0	0	1.17	1	0.854	0	0.905	0
TR00204	44.28	45.16	3.15	6.82	1.91	1.70	0	0	0	1.75	1	0.825	0	0.881	0
TR02801	16.37	17.18	4.56	9.88	1.76	1.64	0	0	0	1.53	1	0.847	0	0.906	0
NR00205	58.6	59.21	3.14	8.20	1.59	1.50	0	0	0	1.75	0	0.752	0	0.831	0
NR00205	52.62	58.6	30.12	7.86	1.56	1.47	0	0	0	1.12	1	0.730	0	0.813	0
TR03302	2.74	5.01	4.53	3.08	1.55	1.39	0	0	0	1.53	1	0.591	0	0.665	0
TR02801	17.18	19.93	16.19	9.09	1.54	1.47	0	0	0	1.18	1	0.745	0	0.829	0
NR00205	40.64	49.34	38.93	6.67	1.49	1.41	0	0	0	1.10	1	0.671	0	0.759	0

Continue ...

Table B2.1 (continued)

Road	Start	End	E	N	R	m	CN1	CR1	NR	R _c	CR2	P _{b1}	B1	P _{b2}	B2
TR00202	37.09	42.36	56.21	15.75	1.48	1.44	1	0	0	1.08	1	0.791	0	0.882	0
TR02801	19.93	23.6	22.73	8.99	1.45	1.40	0	0	0	1.14	1	0.692	0	0.787	0
NR00205	0	2.02	12.44	8.91	1.45	1.39	0	0	0	1.22	1	0.688	0	0.784	0
DR01101	1.9	5.79	11.25	3.86	1.33	1.28	0	0	0	1.24	1	0.534	0	0.619	0
TR00204	2.84	4.4	6.79	5.77	1.32	1.28	0	0	0	1.37	0	0.564	0	0.661	0
NR00105	29.83	31.89	7.86	4.85	1.27	1.24	0	0	0	1.33	0	0.521	0	0.615	0
DR01105	0	3.99	39.16	12.28	1.25	1.24	0	0	0	1.10	1	0.570	0	0.701	0
TR02801	3.73	9.1	23.78	5.40	1.22	1.20	0	0	0	1.14	1	0.494	0	0.594	0
MR00216	0	3.45	22.78	7.83	1.19	1.18	0	0	0	1.14	1	0.484	0	0.601	0
TR00204	45.16	50.54	19.59	4.28	1.17	1.17	0	0	0	1.16	1	0.458	0	0.550	0
TR02801	23.6	25.83	15.65	8.07	1.15	1.15	0	0	0	1.19	0	0.455	0	0.574	0
DR01101	0	1.9	10.49	6.32	1.14	1.15	0	0	0	1.26	0	0.446	0	0.554	0
TR02901	39.92	45.21	8.08	1.70	1.11	1.13	0	0	0	1.32	0	0.414	0	0.483	0
MR00279	33.98	34.65	1.97	2.99	1.01	1.06	0	0	0	2.17	0	0.366	0	0.448	0
TR02901	51.85	55.08	5.00	1.55	1.00	1.07	0	0	0	1.49	0	0.374	0	0.442	0
NR00205	7.12	9.12	18.06	9.00	1.00	1.02	0	0	0	1.17	0	0.302	0	0.424	0
MR00187	6.96	8.03	9.21	8.41	0.98	1.00	0	0	0	1.29	0	0.287	0	0.404	0
MR00279	34.65	35.33	3.12	4.41	0.96	1.00	0	0	0	1.76	0	0.316	0	0.408	0
NR00208	65.71	67.78	16.38	7.25	0.92	0.95	0	0	0	1.18	0	0.241	0	0.346	0
NR00107	0	34.86	128.6	3.36	0.91	0.98	0	0	0	1.06	0	0.298	0	0.380	0
TR03302	1.74	2.74	2.20	2.00	0.91	1.00	0	0	0	2.05	0	0.327	0	0.399	0
TR03201	22.79	26.21	7.81	2.05	0.90	0.99	0	0	0	1.33	0	0.320	0	0.392	0
TR03201	5.71	15.26	21.85	1.99	0.87	0.98	0	0	0	1.15	0	0.308	0	0.379	0
MR00191	16.66	20.2	19.69	4.80	0.86	0.92	0	0	0	1.16	0	0.235	0	0.323	0
TR03201	15.26	22.79	15.09	1.73	0.86	0.98	0	0	0	1.19	0	0.313	0	0.381	0
NR00108	62	72.44	37.60	3.07	0.85	0.93	0	0	0	1.10	0	0.267	0	0.345	0
MR00227	1	5.89	9.71	1.64	0.82	0.96	0	0	0	1.27	0	0.299	0	0.365	0

Continue ...

Table B2.1 (continued)

Road	Start	End	E	N	R	m	CN1	CR1	NR	R _c	CR2	P _{b1}	B1	P _{b2}	B2
TR03102	1.94	9.04	37.92	4.37	0.82	0.89	0	0	0	1.10	0	0.210	0	0.291	0
TR00202	54.11	58.21	21.11	4.15	0.81	0.88	0	0	0	1.15	0	0.206	0	0.286	0
NR00205	9.12	9.85	6.23	6.85	0.80	0.85	0	0	0	1.40	0	0.149	0	0.233	0
NR00108	56.93	62	18.89	2.96	0.79	0.89	0	0	0	1.16	0	0.233	0	0.307	0
MR00027	42.91	44.81	20.18	8.42	0.79	0.83	0	0	0	1.15	0	0.118	0	0.199	0
MR00187	4.36	6.96	24.34	7.31	0.78	0.83	0	0	0	1.14	0	0.124	0	0.203	0
NR00206	35.76	37.22	7.78	4.11	0.77	0.85	0	0	0	1.33	0	0.183	0	0.258	0
MR00027	44.81	50.67	59.26	7.68	0.76	0.81	0	0	0	1.08	0	0.103	0	0.176	0
NR00108	39.85	56.93	56.82	2.52	0.76	0.87	0	0	0	1.08	0	0.225	0	0.294	0
MR00187	8.03	14.71	60.22	6.74	0.75	0.80	0	0	0	1.08	0	0.109	0	0.180	0
TR03103	1.9	18.43	31.01	1.39	0.74	0.91	0	0	0	1.12	0	0.272	0	0.335	0
NR00205	32.72	40.64	39.18	3.66	0.74	0.83	0	0	0	1.10	0	0.174	0	0.244	0
TR02801	9.1	16.37	30.07	3.03	0.73	0.84	0	0	0	1.12	0	0.189	0	0.258	0
TR03102	10.78	13.58	11.14	2.86	0.72	0.83	0	0	0	1.24	0	0.187	0	0.254	0
TR03201	32.73	42.84	23.09	1.58	0.69	0.86	0	0	0	1.14	0	0.236	0	0.297	0
MR00027	68.56	72.13	34.54	6.44	0.67	0.73	0	0	0	1.11	0	0.063	0	0.113	0
NR00106	0.42	29.88	108.2	2.44	0.67	0.80	0	0	0	1.07	0	0.171	0	0.232	0
TR00202	52.64	54.11	9.12	4.08	0.66	0.75	0	0	0	1.29	0	0.106	0	0.163	0
NR00208	0	6.35	68.46	6.93	0.64	0.70	0	0	0	1.08	0	0.045	0	0.087	0
MR00166	3.93	4.71	1.60	1.28	0.62	0.84	0	0	0	2.43	0	0.221	0	0.278	0
MR00188	16.54	19.17	9.76	2.28	0.61	0.76	0	0	0	1.27	0	0.148	0	0.204	0
MR00344	0	6.2	20.35	1.94	0.59	0.75	0	0	0	1.15	0	0.152	0	0.207	0
NR00107	61.64	68.28	27.54	2.41	0.58	0.72	0	0	0	1.13	0	0.120	0	0.172	0
NR00205	2.02	7.12	46.50	5.29	0.58	0.65	0	0	0	1.09	0	0.039	0	0.074	0
TR00203	26.1	38.24	36.61	1.73	0.57	0.75	0	0	0	1.11	0	0.157	0	0.210	0
NR00208	46.62	48.63	15.85	4.48	0.57	0.65	0	0	0	1.18	0	0.048	0	0.084	0
TR02901	24.03	27.33	5.44	0.91	0.55	0.82	0	0	0	1.45	0	0.221	0	0.274	0

Continue ...

Table B2.1 (continued)

Road	Start	End	E	N	R	m	CN1	CR1	NR	R _c	CR2	P _{b1}	B1	P _{b2}	B2
NR00206	25.35	27.67	12.78	3.02	0.55	0.67	0	0	0	1.22	0	0.076	0	0.118	0
NR00108	3.61	39.85	137.7	2.07	0.54	0.70	0	0	0	1.06	0	0.117	0	0.165	0
NR00206	16.61	25.35	34.91	2.17	0.54	0.70	0	0	0	1.11	0	0.111	0	0.158	0
NR00105	81.64	84.46	11.13	2.13	0.54	0.70	0	0	0	1.24	0	0.110	0	0.158	0
NR00205	59.21	66.75	40.83	2.92	0.54	0.66	0	0	0	1.10	0	0.074	0	0.116	0
NR00107	34.86	61.64	104.2	2.09	0.54	0.70	0	0	0	1.07	0	0.111	0	0.158	0
NR00205	70.77	82.05	69.00	3.28	0.54	0.65	0	0	0	1.08	0	0.061	0	0.099	0
NR00206	46.84	58.17	68.10	3.18	0.53	0.64	0	0	0	1.08	0	0.060	0	0.098	0
TR00202	58.21	64.28	41.63	3.62	0.53	0.63	0	0	0	1.10	0	0.048	0	0.082	0
NR00206	58.17	62.43	23.16	2.82	0.52	0.65	0	0	0	1.14	0	0.067	0	0.106	0
NR00206	38.65	45.08	31.25	2.33	0.48	0.63	0	0	0	1.12	0	0.067	0	0.103	0
NR00208	63.69	65.71	17.72	3.96	0.45	0.55	0	0	0	1.17	0	0.016	0	0.032	0
NR00104	27.78	54.24	139.5	2.27	0.43	0.58	0	0	0	1.06	0	0.046	0	0.074	0
NR00105	39.77	81.64	160.0	1.62	0.42	0.62	0	0	0	1.06	0	0.077	0	0.112	0
MR00177	18.12	22.51	48.07	4.56	0.42	0.50	0	0	0	1.09	0	0.006	0	0.013	0
NR00105	31.89	39.77	31.26	1.65	0.42	0.60	0	0	0	1.12	0	0.070	0	0.104	0
DR01105	3.99	5.12	7.36	2.65	0.41	0.54	0	0	0	1.35	0	0.025	0	0.045	0
MR00200	3.64	11.34	83.98	4.42	0.40	0.49	0	0	0	1.07	0	0.005	0	0.012	0
TR07501	29.55	31.49	7.48	1.55	0.40	0.60	0	0	0	1.34	0	0.069	0	0.102	0
MR00191	20.2	20.7	2.61	2.00	0.38	0.54	0	0	0	1.89	0	0.036	0	0.060	0
MR00223	0	6.3	36.96	2.22	0.38	0.52	0	0	0	1.10	0	0.027	0	0.047	0
NR00206	45.08	46.84	8.20	1.70	0.37	0.54	0	0	0	1.32	0	0.042	0	0.067	0
NR00206	62.43	68.3	27.38	1.70	0.37	0.54	0	0	0	1.13	0	0.042	0	0.066	0
NR00206	12.44	16.61	17.07	1.44	0.35	0.55	0	0	0	1.17	0	0.051	0	0.077	0
NR00205	66.75	70.77	20.16	1.74	0.35	0.52	0	0	0	1.15	0	0.033	0	0.053	0
NR00208	48.63	54.91	57.44	2.87	0.31	0.42	0	0	0	1.08	0	0.004	0	0.008	0
NR00208	8.27	16.78	74.11	2.70	0.31	0.42	0	0	0	1.07	0	0.004	0	0.010	0

Continue ...

Table B2.1 (continued)

Road	Start	End	E	N	R	m	CN1	CR1	NR	R _c	CR2	P _{b1}	B1	P _{b2}	B2
NR00206	68.3	72.85	27.20	1.54	0.26	0.42	0	0	0	1.13	0	0.011	0	0.020	0
NR00205	11.4	14.67	24.35	1.83	0.25	0.39	0	0	0	1.14	0	0.004	0	0.009	0
NR00207	43.29	52.07	95.12	2.62	0.24	0.34	0	0	0	1.07	0	0.001	0	0.002	0
NR00208	59.42	63.69	37.50	1.87	0.21	0.34	0	0	0	1.10	0	0.001	0	0.003	0
NR00208	34.32	46.62	109.1	1.87	0.21	0.34	0	0	0	1.07	0	0.001	0	0.003	0
MR00199	19.57	22.46	15.84	1.04	0.19	0.38	0	0	0	1.18	0	0.009	0	0.016	0

APPENDIX B3

COMPARISON OF RANKING METHODS

The 113 segments identified in Appendix B1 were ranked according to the following criteria :

- Accident number - X (acc/km)
- Accident rate – R (acc/mvkm)
- Bayesian safety estimate - m (acc/mvkm)
- Potential Accident Reduction - PAR (acc/year)

The results of this ranking study are shown in Table B3.1.

Table B3.1 : Comparison of ranking procedures

Road	Start	End	AADT	E	N	Rank(N)	R	Rank(R)	M	Rank(m)	PAR	Rank(PAR)
MR00165	3.63	7.47	5530	31.02	19.8	4	2.45	12	2.40	10	13.23	1
TR00202	37.09	42.36	7300	56.21	15.7	7	1.48	27	1.47	26	10.30	2
NR00205	40.64	49.34	3063	38.93	6.7	33	1.49	26	1.48	24	7.26	3
MR00165	0	3.63	1502	7.97	8.8	17	4.02	3	3.66	2	6.52	4
NR00205	52.62	58.6	3447	30.12	7.9	23	1.56	23	1.55	22	6.15	5
NR00107	0	34.86	2526	128.65	3.4	61	0.91	46	0.91	48	5.32	6
TR03201	0	5.71	1729	14.42	5.6	40	2.22	15	2.14	14	5.32	7
NR00205	49.34	51.8	4107	14.76	13.0	8	2.17	17	2.10	15	5.25	8
TR00204	50.54	55.03	2663	17.47	7.6	26	1.95	19	1.90	18	5.25	9
NR00205	51.88	52.62	2704	2.92	31.1	1	7.87	1	5.98	1	5.21	10
DR01105	0	3.99	6717	39.16	12.3	9	1.25	33	1.25	33	4.97	11
NR00205	9.85	11.4	5593	12.67	18.7	5	2.29	14	2.20	13	4.89	12
MR00027	67.19	68.56	4722	9.45	18.2	6	2.65	10	2.48	8	4.49	13
MR00227	5.89	9.66	1111	6.12	5.6	41	3.43	4	3.07	4	4.11	14
TR02801	19.93	23.6	4240	22.73	9.0	15	1.45	28	1.44	28	4.02	15
MR00223	6.3	9.63	1618	7.87	6.3	37	2.67	8	2.48	9	3.79	16
TR02801	17.18	19.93	4029	16.19	9.1	13	1.54	25	1.52	23	3.24	17
MR00027	51.15	51.73	7077	6.00	29.3	2	2.83	7	2.57	7	3.13	18
TR02801	2.14	3.73	3762	8.74	11.9	10	2.17	16	2.06	16	3.12	19
MR00027	51.73	52.29	4220	3.45	26.8	3	4.34	2	3.55	3	3.11	20
TR02801	3.73	9.1	3031	23.78	5.4	42	1.22	34	1.22	34	2.83	21
MR00216	0	3.45	4520	22.78	7.8	24	1.19	35	1.18	35	2.51	22
TR02801	0	2.14	1158	3.62	5.6	39	3.31	6	2.80	5	2.33	23
NR00205	0	2.02	4214	12.44	8.9	16	1.45	29	1.42	29	2.19	24
TR00204	45.16	50.54	2492	19.59	4.3	52	1.17	36	1.17	36	2.11	25
TR03201	42.84	44.35	1708	3.77	6.6	34	2.65	9	2.31	11	1.80	26

Continue ...

Table B3.1 (continued)

Road	Start	End	AADT	E	N	Rank(N)	R	Rank(R)	M	Rank(m)	PAR	Rank(PAR)
DR01101	1.9	5.79	1980	11.25	3.9	57	1.33	30	1.32	30	1.66	27
NR00108	2.68	3.61	2530	3.44	9.7	12	2.62	11	2.25	12	1.61	28
TR02801	23.6	25.83	4803	15.65	8.1	22	1.15	37	1.15	37	1.59	29
TR03201	44.35	45.15	1760	2.06	8.8	18	3.40	5	2.60	6	1.37	30
MR00313	1.6	3.1	1843	4.04	5.3	43	1.98	18	1.80	19	1.25	31
TR02801	16.37	17.18	3850	4.56	9.9	11	1.76	21	1.64	21	1.15	32
NR00205	7.12	9.12	6182	18.06	9.0	14	1.00	42	1.01	43	1.14	33
DR01101	0	1.9	3779	10.49	6.3	36	1.14	38	1.14	38	1.05	34
NR00105	29.83	31.89	2610	7.86	4.9	45	1.27	32	1.26	32	1.04	35
NR00108	62	72.44	2465	37.60	3.1	65	0.85	52	0.86	52	1.01	36
TR00204	2.84	4.4	2981	6.79	5.8	38	1.32	31	1.30	31	0.99	37
TR00204	44.28	45.16	2449	3.15	6.8	31	1.91	20	1.71	20	0.91	38
TR03302	2.74	5.01	1365	4.53	3.1	64	1.55	24	1.47	27	0.91	39
DR01056	0	1.32	1108	2.14	3.8	58	2.34	13	1.93	17	0.85	40
TR02901	39.92	45.21	1045	8.08	1.7	101	1.11	39	1.12	39	0.75	41
NR00208	65.71	67.78	5417	16.38	7.2	28	0.92	45	0.93	46	0.70	42
TR03102	1.94	9.04	3656	37.92	4.4	51	0.82	54	0.83	56	0.70	43
TR03201	5.71	15.26	1566	21.85	2.0	91	0.87	49	0.88	49	0.69	44
NR00205	58.6	59.21	3519	3.14	8.2	21	1.59	22	1.48	25	0.67	45
MR00191	16.66	20.2	3808	19.69	4.8	46	0.86	50	0.88	51	0.59	46
MR00187	6.96	8.03	5894	9.21	8.4	20	0.98	43	1.00	44	0.54	47
TR03201	15.26	22.79	1372	15.09	1.7	98	0.86	51	0.88	50	0.44	48
TR00202	54.11	58.21	3524	21.11	4.1	53	0.81	55	0.82	57	0.32	49
TR02901	51.85	55.08	1060	5.00	1.5	106	1.00	41	1.03	41	0.32	50
TR03201	22.79	26.21	1564	7.81	2.0	88	0.90	48	0.93	47	0.30	51
MR00027	42.91	44.81	7270	20.18	8.4	19	0.79	58	0.81	60	0.25	52
NR00108	56.93	62	2550	18.89	3.0	69	0.79	57	0.81	59	0.24	53

Continue ...

Table B3.1 (continued)

Road	Start	End	AADT	E	N	Rank(N)	R	Rank(R)	M	Rank(m)	PAR	Rank(PAR)
MR00027	44.81	50.67	6922	59.26	7.7	25	0.76	61	0.77	62	0.23	54
MR00187	4.36	6.96	6407	24.34	7.3	27	0.78	59	0.80	61	0.22	55
MR00227	1	5.89	1359	9.71	1.6	103	0.82	53	0.86	53	0.19	56
NR00108	39.85	56.93	2277	56.82	2.5	77	0.76	62	0.76	63	0.18	57
MR00279	34.65	35.33	3142	3.12	4.4	50	0.96	44	1.01	42	0.17	58
MR00279	33.98	34.65	2016	1.97	3.0	68	1.01	40	1.06	40	0.13	59
NR00205	9.12	9.85	5839	6.23	6.8	30	0.80	56	0.86	54	0.09	60
TR03302	1.74	2.74	1507	2.20	2.0	89	0.91	47	0.99	45	0.09	61
NR00206	35.76	37.22	3647	7.78	4.1	54	0.77	60	0.82	58	0.05	62
MR00187	8.03	14.71	6170	60.22	6.7	32	0.75	63	0.75	66	0.05	63
TR03103	1.9	18.43	1284	31.01	1.4	110	0.74	64	0.76	65	-0.02	64
NR00205	32.72	40.64	3386	39.18	3.7	59	0.74	65	0.75	67	-0.04	65
MR00166	3.93	4.71	1405	1.60	1.3	111	0.62	73	0.84	55	-0.05	66
TR03102	10.78	13.58	2722	11.14	2.9	72	0.72	67	0.76	64	-0.07	67
TR02801	9.1	16.37	2831	30.07	3.0	66	0.73	66	0.75	68	-0.09	68
TR00202	52.64	54.11	4247	9.12	4.1	55	0.66	71	0.71	70	-0.20	69
MR00191	20.2	20.7	3576	2.61	2.0	90	0.38	100	0.62	77	-0.24	70
TR02901	24.03	27.33	1129	5.44	0.9	113	0.55	80	0.66	74	-0.26	71
TR03201	32.73	42.84	1563	23.09	1.6	105	0.69	68	0.71	69	-0.29	72
MR00188	16.54	19.17	2540	9.76	2.3	81	0.61	74	0.67	72	-0.32	73
NR00105	81.64	84.46	2701	11.13	2.1	85	0.54	84	0.60	81	-0.57	74
DR01105	3.99	5.12	4461	7.36	2.7	75	0.41	97	0.51	92	-0.62	75
NR00206	25.35	27.67	3770	12.78	3.0	67	0.55	81	0.60	80	-0.63	76
TR07501	29.55	31.49	2640	7.48	1.5	107	0.40	99	0.50	94	-0.64	77
MR00027	68.56	72.13	6622	34.54	6.4	35	0.67	69	0.68	71	-0.67	78
NR00208	46.62	48.63	5398	15.85	4.5	48	0.57	79	0.61	78	-0.70	79
NR00206	45.08	46.84	3189	8.20	1.7	99	0.37	102	0.46	96	-0.78	80

Continue ...

Table B3.1 (continued)

Road	Start	End	AADT	E	N	Rank(N)	R	Rank(R)	M	Rank(m)	PAR	Rank(PAR)
MR00344	0	6.2	2247	20.35	1.9	92	0.59	75	0.62	76	-0.79	81
NR00107	61.64	68.28	2839	27.54	2.4	79	0.58	76	0.60	79	-1.12	82
NR00208	63.69	65.71	6003	17.72	4.0	56	0.45	92	0.49	95	-1.30	83
NR00206	58.17	62.43	3721	23.16	2.8	73	0.52	90	0.55	87	-1.31	84
TR00203	26.1	38.24	2064	36.61	1.7	97	0.57	78	0.59	83	-1.56	85
NR00206	12.44	16.61	2802	17.07	1.4	109	0.35	104	0.40	103	-1.68	86
NR00208	0	6.35	7379	68.46	6.9	29	0.64	72	0.65	75	-1.73	87
NR00206	16.61	25.35	2734	34.91	2.2	84	0.54	83	0.56	84	-1.74	88
NR00205	2.02	7.12	6241	46.50	5.3	44	0.58	77	0.59	82	-1.90	89
NR00205	66.75	70.77	3432	20.16	1.7	96	0.35	105	0.39	105	-2.00	90
NR00206	38.65	45.08	3327	31.25	2.3	80	0.48	91	0.50	93	-2.06	91
NR00205	59.21	66.75	3706	40.83	2.9	70	0.54	85	0.56	85	-2.09	92
NR00106	0.42	29.88	2514	108.21	2.4	78	0.67	70	0.67	73	-2.13	93
MR00199	19.57	22.46	3751	15.84	1.0	112	0.19	113	0.25	110	-2.20	94
TR00202	58.21	64.28	4694	41.63	3.6	60	0.53	89	0.55	89	-2.24	95
NR00105	31.89	39.77	2715	31.26	1.6	102	0.42	96	0.44	97	-2.56	96
NR00206	62.43	68.3	3193	27.38	1.7	100	0.37	103	0.40	104	-2.59	97
NR00205	11.4	14.67	5096	24.35	1.8	95	0.25	109	0.29	109	-3.03	98
NR00206	68.3	72.85	4092	27.20	1.5	108	0.26	108	0.29	108	-3.31	99
MR00223	0	6.3	4015	36.96	2.2	83	0.38	101	0.40	102	-3.37	100
NR00205	70.77	82.05	4187	69.00	3.3	62	0.54	87	0.55	88	-3.58	101
NR00206	46.84	58.17	4114	68.10	3.2	63	0.53	88	0.54	91	-3.67	102
MR00177	18.12	22.51	7495	48.07	4.6	47	0.42	95	0.43	99	-3.94	103
NR00208	59.42	63.69	6011	37.50	1.9	93	0.21	111	0.24	112	-4.97	104
NR00107	34.86	61.64	2665	104.27	2.1	86	0.54	86	0.54	90	-5.39	105
NR00208	48.63	54.91	6260	57.44	2.9	71	0.31	106	0.33	106	-6.18	106
NR00108	3.61	39.85	2602	137.77	2.1	87	0.54	82	0.55	86	-6.87	107

Continue ...

Table B3.1 (continued)

Road	Start	End	AADT	E	N	Rank(N)	R	Rank(R)	M	Rank(m)	PAR	Rank(PAR)
MR00200	3.64	11.34	7465	83.98	4.4	49	0.40	98	0.41	101	-7.12	108
NR00208	8.27	16.78	5961	74.11	2.7	74	0.31	107	0.32	107	-8.03	109
NR00104	27.78	54.24	3609	139.52	2.3	82	0.43	93	0.44	98	-10.95	110
NR00207	43.29	52.07	7415	95.12	2.6	76	0.24	110	0.25	111	-11.94	111
NR00105	39.77	81.64	2617	160.09	1.6	104	0.42	94	0.43	100	-12.78	112
NR00208	34.32	46.62	6074	109.15	1.9	94	0.21	112	0.22	113	-14.55	113