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24

25 **Abstract** (228 words)

26 Ecosystem-scale examination of fish communities typically involves creating  
27 spatiotemporally-explicit relative abundance distribution maps using data derived from  
28 multiple fishery-independent surveys. However, survey sampling performance varies by  
29 vessel and sampling gear, which may influence estimated species distribution patterns. Using  
30 generalised additive mixed models, the effect of different gear-vessel combinations on  
31 relative abundance estimates at length are investigated using European fisheries-independent  
32 groundfish survey data. We constructed a modelling framework for evaluating relative  
33 efficiency of multiple survey gear-vessel combinations and examined 19 disparate surveys  
34 for 254 species-length combinations across the northeast Atlantic. Space-time variables  
35 explained the majority of the variation in catches when combining data across different gears  
36 or vessels for 181 of 254 species-length cases, indicating that for many species, models could  
37 successfully characterize distribution patterns by combining data from disparate surveys.  
38 Variables controlling for catch efficiency differences across gear-vessel combinations  
39 explained substantial variation in catches for 127 of 254 species-length data sets. In such  
40 cases, models that fail to control for gear efficiencies across surveys can mask changes in the  
41 spatial distribution of species. Estimated relative differences in catch efficiencies grouped  
42 strongly by gear type, but did not exhibit a clear pattern across species' functional forms,  
43 suggesting difficulty in predicting the potential impact of gear efficiency differences when  
44 combining data across surveys to assess species' distributions and highlighting the  
45 importance of modelling approaches that can control for gear differences.

46

47 **Keywords**

48 Catchability; gear efficiency; fisheries independent assessment; Generalised Additive Mixed  
49 Model (GAMM); survey standardisation; species distribution modelling

50 **1. Introduction**

51 As ecosystem-based management in the marine environment advances, fisheries policies  
52 increasingly require consideration of both target and non-target species in assessing the state  
53 of fisheries and fishing impacts on marine ecosystems (e.g. the European Union (EU) Marine  
54 Strategy Framework Directive (MSFD; EC 2008; 2010; 2017), Common Fisheries Policy  
55 (CFP; EC, 2013), United States Magnuson–Stevens Fishery Conservation and Management  
56 Act (US, 1996; 2006), etc.). This transition to ecosystem-based management has led to a need  
57 for greater understanding and detailed information on the distribution of a broad spectrum of  
58 fish species across large spatial scales, such as large marine ecosystems or ecoregions (Kelley  
59 & Sherman 2018).

60

61 Fisheries-independent groundfish surveys sample both commercial and non-target fish  
62 species, often providing the only data source available to estimate relative abundances for  
63 non-commercial species (Poos et al., 2013). These surveys tend to be discrete monitoring  
64 programmes, operating at local scales usually associated with the exclusive economic zones  
65 of countries managing the surveys. To obtain information on fish distributions at large marine  
66 ecosystems scales, therefore, requires integration across national jurisdictional boundaries  
67 and multiple disparate surveys that may differ in terms of spatial coverage, survey vessel,  
68 season, types of fishing gear, and survey protocols. Amalgamating such data into a single  
69 cohesive analysis is difficult because of potential differences in gear efficiency among  
70 different length-classes and species of fish (Fraser et al., 2007; Walker et al., 2017), types of  
71 survey gear, and vessels that vary in their fishing power (Dann et al., 2005).

72

73 Estimates of species' latent abundance, and hence species-at-length catchability coefficients,  
74 are rarely available in fisheries survey data. In isolation, each individual survey provides

75 estimates of species' relative abundance at sampled locations and can provide assessment of  
76 the spatial distribution of fish within the survey domain. Problems may arise, however, when  
77 two or more surveys need to be combined to assess species' distributions. If gear efficiencies  
78 vary between different surveys, then estimates of species relative abundance provided by  
79 each survey may not be compatible. Failure to understand, or ignoring, how gear efficiency  
80 differs between surveys may lead to incoherent abundance estimates when merging surveys  
81 together to conduct assessments at large spatial scales. To perform such assessment,  
82 therefore, requires quantification of gear efficiency for different species, different size classes  
83 of fish, and different gears.

84

85 The traditional approach to estimating gear efficiency is through paired field experiments,  
86 where two vessels fish side by side and compare catches (Somerton et al., 1999; Zhou et al.,  
87 2014). Such experiments are costly to conduct and are generally implemented over limited  
88 spatial and temporal scales. However, where different survey domains overlap spatially, there  
89 may be opportunity to utilize species distribution modelling to complement, or even replace,  
90 field-based estimation of gear efficiencies (e.g. Ono et al., 2018); thereby providing a  
91 convenient framework for handling data from disparate surveys that can be regularly updated  
92 as new survey data become available. Statistical modelling of species distributions from large  
93 data sets is no longer limited by insufficient computing capacity. The use of such models  
94 offers an opportunity of overcoming challenges in combining data across surveys with  
95 varying gear efficiencies to enable extensive study of marine species distributions across  
96 large spatial scales.

97

98 Here we build from previous gear efficiency modelling efforts (Walker et al., 2017, Zhou et  
99 al., 2014) with an aim to advance the tools available for combining information across

100 disparate fisheries surveys towards informing the spatial ecology of marine species. The  
101 spatial scale, the number of species assessed, the interaction between the gear-vessel  
102 combinations, and the spatial and temporal variation inherent within European fisheries  
103 surveys presents unique challenges requiring a new approach. Utilizing Generalised Additive  
104 Mixed Models (GAMMs); we analyse the proportion of variance explained by the differences  
105 in gear efficiency and the spatial–temporal variation in abundance of 135 species, in three  
106 length categories, collected in the 19 northeast Atlantic groundfish surveys with 24 different  
107 gear-vessel combinations. Here we focus on bottom trawl gears, namely otter trawls and  
108 beam trawls, as others have previously focused on combining acoustic measurements with  
109 habitat data to gain inference about the abundance of fish and infer on bottom trawl gear  
110 efficiencies (Kotwicki et al., 2018). Three length categories were chosen to (1) capture the  
111 main intra-specific length-related catchability differences described in previous studies  
112 (Fraser et al., 2007; Walker et al., 2017), (2) broadly reflect trophic guilds in marine fish  
113 communities (ICES, 2017), and (3) reflect the main size classes of fish either retained in  
114 commercial trawls or that escape through the mesh (Piet et al., 2009). The 24 gear-vessel  
115 combinations were chosen to best reflect the perceived differences in rigging and standard  
116 operating procedures carried out by different countries in their national surveys (Table 1). By  
117 understanding which species in our length categories are affected by variations among gears  
118 and vessels, our primary goal is to develop a consistent approach for combining groundfish  
119 surveys to facilitate marine ecosystem monitoring at large spatial scales. Using the GAMMs  
120 to control for differences in gear efficiency among surveys, we also generate estimates of  
121 spatial and temporal trends of relative abundance for species among different length  
122 categories throughout the northeast Atlantic to inform marine fish community ecological  
123 analyses (covering three ICES marine ecoregions/large marine ecosystems: Greater North  
124 Sea, Celtic Seas, and Bay of Biscay and the Iberian Coastal; Spalding et al., 2007). Finally,

125 we conclude with a discussion of high priority information needs to further improve  
126 understanding of gear efficiency within marine fisheries survey data.

127

## 128 **2. Methods**

### 129 *2.1 Fisheries Surveys*

130 Data for most European groundfish surveys are uploaded and maintained on the ICES  
131 “Database of Trawl Surveys” (DATRAS). Data for surveys carried out in the Northeast  
132 Atlantic were recently subjected to a quality assurance and quality audit (QAQA) process  
133 (Moriarty et al., 2017; Greenstreet and Moriarty 2017a; 2017b; Moriarty et al., 2019), to  
134 ensure their adequacy to meet monitoring and assessment requirements under the EU MSFD  
135 (EC, 2008; 2010; 2017). These standard monitoring programme data products, along with  
136 data for four Spanish surveys, which underwent the same QAQA process but were not fully  
137 uploaded to DATRAS, were used in this study to obtain maximum spatial and temporal  
138 coverage and include the widest possible range of survey types for modelling (Table 1). Each  
139 survey data product includes the number of fish caught ( $C_{i,s,l}$ ) of a species ( $s$ ) at length ( $l$ ),  
140 for each trawl sample ( $i$ ), along with the vessel and fishing gear ( $g$ ), tow location, date,  
141 depth and swept area ( $E$ ). The fishing gear ( $g$ ), included information from vessels that were  
142 expected to fish differently based on their gear configuration information. For example, both  
143 French and Irish vessels surveying in the Celtic Seas region use a GOV gear. However, the  
144 French surveys use double sweeps, and the Irish surveys rotate between a standard GOV  
145 survey gear (ICES 2015) and a double sweep with 16-inch bobbins, depending on the  
146 substrate (Table 1). The fish abundance data were organized into three broad length  
147 categories ( $lc$ ), small unfished (<23cm), intermediate transition (23 - 35cm), and large fished  
148 (>35cm). Groundfish surveys only record those species and lengths caught (i.e. presence only  
149 data). Data rows for zero catches were added to the full data set where species at length were

150 not reported in any given sample. To ensure constant and equivalent distance units, survey  
151 sample latitude - longitude coordinates were converted to eastings and northings ( $X, Y$ ) using  
152 R package “Rgdal” (Bivand et al., 2018). Date ( $t$ ) was incremented in quarterly time bins  
153 starting from quarter 4 (Oct – Dec) 2003, which was assigned time step  $t = 1$ , while the  
154 quarter 1 (Jan – Mar) 2004 was assigned time step  $t = 2$ , and so on.

155

## 156 *2.2 Exploring Sources of Variation in Survey Abundance at Length Data*

157 Generalised Additive Mixed Models (GAMMs) were used to account for non-linear spatial  
158 and temporal trends in fish density while simultaneously estimating gear efficiency using a  
159 modelling framework adapted from Walker et al. (2017). Survey catches were modelled as  
160 counts, with separate regressions for each species-length bin combination. Many species had  
161 a preponderance of zero catches. Initial exploration casting GAMMs for all species within  
162 Poisson, negative binomial, and zero-inflated Poisson frameworks showed that Poisson  
163 models provided a poor fit and failed to accommodate over-dispersion in catch data. Negative  
164 binomial and zero-inflated Poisson models showed similar fits for non-schooling species, but  
165 schooling species violated the assumption of independence required by Poisson processes.  
166 Consequently, we analyzed catches as Negative Binomially (NB) distributed GAMMs fit  
167 using the “mgcv” package (Wood 2004; 2011) in the R statistical programming environment  
168 (R Core Team 2017). The full model for a given species and length category catch data set  
169 had the form:

$$170 C_i \sim NB(\mu_i, k)$$

$$171 \text{ with } E[C_i] = \mu_i = e^{\log(E_i) + s(X_i, Y_i, t_i) + zg_{(i)}} \quad 1,$$

172 where  $C_i$  is the number of fish of a given species in a given length category caught in the  $i^{\text{th}}$   
173 sample (fishing event),  $k$  is the negative binomial shape parameter representing the degree of

174 overdispersion,  $\log(E_i)$  is the log of swept area for fishing event  $i$  which was included as an  
175 offset to account for varying fishing effort among trips,  $s(X_i, Y_i, t_i)$  denotes a multivariate  
176 smoothing function to represent spatio-temporal trends in catch data, and  $z_{g(i)}$  are i.i.d.  
177 normally distributed random effects for gear-vessel combinations associated with fishing  
178 events. The space-time smoothing model component,  $s(X_i, Y_i, t_i)$ , was specified as a tensor  
179 product smoother for which the associated basis functions were cast as cubic splines with  
180 shrinkage (i.e.,  $te(X_i, Y_i, t_i, bs = "cs")$  in mgcv formulaic notation), a formulation which can  
181 accommodate data on different scales (Wood 2004; 2011). Gear-vessel combination was  
182 treated as a random effect, as opposed to a fixed effect, because variation among catch  
183 efficiencies is the primary feature of interest, and because this approach also aids in model  
184 convergence by reducing the number of fitted parameters. The spatiotemporal smoother  
185 describes the underlying estimated distribution of species across space and time; whereas the  
186 random effect controls for variation among gear efficiency when combining disparate survey  
187 data sets. To facilitate model convergence, we excluded data on species-at-length for which  
188 any given length category was sampled by fewer than two gear-vessel combinations or was  
189 sampled fewer than 100 times. The full model was compared to a reduced model that  
190 included space-time covariates, but which did not account for the effect of gear-vessel  
191 combinations (i.e. the gear-vessel combination random effect was dropped) in order to assess  
192 the impact on species distribution modelling inference when gear is ignored. Comparisons of  
193 full and reduced model fits were assessed using Akaike's information criterion (AIC). The  
194 full model was further assessed for reliability using visual tests and a chi squared goodness of  
195 fit test. To substantiate that our GAMM models can effectively differentiate between the  
196 random gear-vessel effects and the spatial and temporal variation in the abundance of  
197 demersal fish in the north east Atlantic region, we performed a simulation-estimation  
198 experiment (Supplemental Material S2).



199

200        *2.3 Interpretation of models*

201    To interpret the importance of gear efficiency versus spatiotemporal distribution patterns in  
202    explaining variation in survey data, we utilized variance components analysis. This analysis  
203    partitions total variation in the fitted data among the three modelled components: gear  
204    efficiency, spatiotemporal distribution, or unexplained residual variation. Accordingly, when  
205    the gear component constitutes the preponderance of model variation for a given species and/  
206    or length category, we conclude that gear efficiency varies widely across gears and surveys.  
207    In contrast, when location and time make up the majority of model variability for a given  
208    species, we conclude that catches are more strongly influenced by the ecology of the fish,  
209    rather than the differences in gear efficiency.

210

211    A non-metric multidimensional scaling (nMDS) unconstrained ordination technique using  
212    Euclidean distances was employed to explore how each species within the assemblages  
213    varied with estimates of gear efficiencies among gear-vessel coefficients and length classes  
214    from our models. Species were grouped by taxonomic order as a proxy for functional forms  
215    to examine if there was a pattern in estimates of gear efficiencies in species groups with  
216    similar morphological or ecological attributes. The gear-vessel coefficients were conditioned  
217    into a matrix, where the Scottish vessel with a GOV gear type was used as a reference gear,  
218    and the difference was calculated for each other gear-vessel combination. Permutational  
219    multivariate analysis of variance (PERMANOVA) was used to test the differences between  
220    the gear-vessel coefficients derived for each species in each length class from our full models  
221    for similar gear types. A clustering criterion that minimizes the amount of variance within in  
222    the gear-vessel groups was implemented (Ward, 1963). Euclidean distance was used and the

223 *p*-value was set to 0.05. The nMDS and PERMANOVA routines were implemented in R (R  
224 Core Team 2017) using the “vegan” package (Oksanen et al., 2017).

225

### 226 **3. Results**

227 Data for 135 fish species were available from otter trawl surveys across the northeast  
228 Atlantic, whereas beam trawl surveys operate in a much more limited area within the North  
229 and Irish Seas (Figure 1). The surveys carried out in the Irish Sea have the highest degree of  
230 spatial and temporal overlap, whereas survey overlap is more limited in the Bay of Biscay  
231 and Iberian Coast region (Figure 1).

232

233 Two hundred and fifty-four full GAMMs were fit to 132 species in up to three length  
234 categories (Figure 2). For fishes in the smallest size class (<23cm), the full model was fit to  
235 109 species, and 23 species had insufficient data based on the criteria described in Methods  
236 (Section 2.2). For fishes in the intermediate transition category (23-35cm), the full model was  
237 fit to 85 species, and 47 species had insufficient data. For the largest size class (>35cm), the  
238 full model was fit to 60 species, and 72 species had insufficient data.

239

240 In 39/254 models, the unexplained variance was greater than the explained variance (Figure  
241 2). In 237/254 of the species-length combinations, the full model, which controlled for  
242 differences in gear-vessel combinations, improved the deviance explained over the reduced  
243 model (Table S1.1). 250/254 full models had a lower AIC score than the reduced model. In  
244 the cases where the full estimates did not improve inference, the differences in the amounts of  
245 deviance explained and the AIC scores between the full and reduced models were small  
246 (Table S1.1).

247

248 In 215/254 full models, over 50% of the variation in the data can be explained, suggesting  
249 that this framework is an effective way of calculating variance in latent species abundance  
250 over a large spatial scale. In 181/254 full models, location ( $X, Y$ ) and time ( $t$ ) components of  
251 the model explained over 50% of the variation in the data, suggesting that catch rates are  
252 strongly driven by the ecology of the fish, while the random effect of fishing gear on a given  
253 vessel( $g$ ) at a given length category ( $l$ ) generally plays a smaller role in explaining variance.  
254 Indeed, in 51 of these 181 models, the overall variance explained is >50%, but the variance  
255 explained by gear is <1%. As an example, for common dab (*Limanda limanda*) in the <23cm  
256 length class, the random effect of fishing gear on a given vessel ( $g$ ) explains 0.007% of the  
257 variance, while location ( $X, Y$ ) and time ( $t$ ) components explained 62.2% of the variance  
258 (Figure 3a/b). In this case, the reduced model, where location ( $X, Y$ ) and time ( $t$ ) components  
259 explained 61.1% of the variance, performed similarly to the full model (Supplemental  
260 Material 1 Table S1.1).

261

262 In 37/254 full models, the overall variance explained is >50%, and the gear component  
263 explains between 1% and 5% of the variation, suggesting that gear efficiency varies across  
264 gears and vessel combinations but has relatively little influence on catch performance. For  
265 example, for the thorny skate (*Amblyraja radiata*) in the 23-35cm length class, the random  
266 effect of fishing gear ( $g$ ) explained 3.7% of the variance, while location ( $X, Y$ ) and time  
267 ( $t$ ) components of the full model explained 68.7% of the variance. While the estimated  
268 variance component for gear effects was smaller than the space-time components, the effect  
269 of fishing gear can be seen in the difference in spatial pattern between the full and reduced  
270 models (Figure 3d).

271

272 In 127/254 full models the overall variance explained is >50%, and the gear component  
273 explains more than 5% of the variation, suggesting that gear efficiency for these species-at-  
274 length varies substantially across gear and vessel combinations. For example, for sole (*Solea*  
275 *solea*) in the 23-35cm length class, the random effect of fishing gear ( $g$ ) explained 8.6% of  
276 the variance, while location ( $X, Y$ ) and time ( $t$ ) components of the full model explained  
277 46.5% of the variance in the data (Figure 3e/f). In this case, the output of the full model  
278 highlights the importance of understanding the effect of fishing gear in assessing the  
279 distribution of this species.

280

281 To assess the difference in inference gleaned from the full and reduced models, we further  
282 explored the spatial-temporal pattern of sole (*Solea solea*) in the 23-35cm category. While the  
283 general pattern is similar in the full and reduced models (Figure 4), the reduced model  
284 suggests the presence of intermediate-sized sole off of the coasts of Spain and Portugal;  
285 whereas the full model suggests that there are no intermediate-sized sole in these areas. When  
286 examined more specifically, we see that for the entire area, the sole data is 88% zero values,  
287 but for the southern part of the study area, where Spain and Portugal survey, the sole data is  
288 96.5% zero values. Consequently, we can conclude that the reduced model is likely to  
289 overestimate the abundance in this area, and that this overestimation is likely an artefact of  
290 not accounting for gear.

291

292 Aggregating over the entire distribution of sole, there is a steadier rate of movement in the  
293 centre of mass in the population estimated from the full model, while the movement in the  
294 centre of mass in the population estimated from the reduced model is more variable (Figure  
295 5a). The centre of mass metric highlights the eastward movement in the population in the full  
296 model, which is not the case in the reduced model (Figure 5b). The inference from the

297 simulations suggests that the full model should be more capable of capturing the direction of  
298 movement than the reduced model ((Supplemental 2, Figure S2.4).

299

300 Unsurprisingly, nMDS highlights that the estimated gear coefficients vary considerably by  
301 gear types (Figure 6a; PERMANOVA test for differences in gears:  $F = 2.36, R^2 = 0.18, p -$   
302  $value = 0.001$ ). However, gear coefficients are largely consistent within gear type,  
303 indicating stable catch efficiencies within gear types regardless of the survey country of  
304 origin or vessel. The GOV, beam trawls, and baca trawls gear-vessels tended to group most  
305 closely in their estimated gear coefficients, whereas other gears tended to differ more widely.  
306 The GOV has the highest level of variance and is the most widely used gear within the  
307 region. The beam trawl surveys have a high level of spatial overlap with the surveys that use  
308 the GOV gear in the North Sea and the rockhopper trawl in the Irish Seas. The baca trawl has  
309 very limited spatial overlap with other gears as it is used exclusively by the Spanish in the  
310 Bay of Biscay and Iberian Coast region. There is no clear pattern emerging in the estimated  
311 relative difference in catch efficiencies across species functional form (Figure 6b).

312

#### 313 **4. Discussion**

314 Understanding how gear efficiency impacts fishery independent survey sampling is required  
315 for robust multi-survey species distribution modelling of both commercial and non-  
316 commercial species and is a key factor in determining absolute abundance estimates for  
317 commercial stocks (Kasatkina & Ivanova, 2009; Maunder & Piner, 2014). The aim of the  
318 analyses presented here is to provide an overall understanding how species are affected by the  
319 rigging of individual vessels to guide future ecosystem-scale species distribution modelling  
320 and examinations of fish communities. Our models support the derivation of relative species  
321 abundance estimates, and they provide information on gear efficiency of 24 gear-vessel

322 combinations seasonally for three length groups chosen to reflect the main intra-specific  
323 length-related differences described in previous catchability studies (Fraser et al., 2007) in  
324 this region. This provides a modelling workflow to combine data across surveys that controls  
325 for potential gear-vessel-specific differences in catchability. The flexible framework  
326 provided here may be adapted to the end users' needs; for example, different length  
327 categories may be applied to answer specific ecological questions. We caution; however, that  
328 the gear efficiency coefficients used in this analysis were estimated using a 10-year historical  
329 time span and are only valid under the conditions for which they are calculated. As such, any  
330 efforts to employ them for correcting individual survey-species catches need take this into  
331 account (Arreguín-Sánchez, 1996).

332 In 15% (39/254) of models, the unexplained variance is higher than the explained variance  
333 (Figure 2). Given that it is unlikely for a species to be randomly distributed in space and time,  
334 this high unexplained variance is likely due to the rareness of the species within a given  
335 length category (i.e. there are not enough samples to describe the latent species distribution).  
336 Species that are rarely caught may not be rare in the environment, but instead may be  
337 particularly poorly sampled (i.e. low gear efficiency) in the survey trawl gear. Sampling of  
338 fish in the marine environment by fishing gear is known to be imperfect (Fraser et al., 2007,  
339 Zhou et al., 2014, Walker et al., 2017). This means additional considerations may need to be  
340 addressed during sampling and data analysis, such as joint dynamic species distribution  
341 modelling (Thorson et al., 2016). Reliable inference depends on sampling methods that  
342 produce reasonable odds of detection given presence, where no estimator will be particularly  
343 helpful when applied to data on populations or species that are “invisible” to collection gear  
344 (MacKenzie et al., 2006).

345

346 The estimated variance components from our models show that in 35% of cases (88/254), the  
347 location and time components explained most of the variation in the data, while the gear  
348 component explained relatively little variation ( $\leq 5\%$ ; Figure 2). This suggests that in such  
349 circumstances, the spatial-temporal distribution of these species can be estimated using  
350 combined survey data. Where the modelled gear component is especially small, particularly  
351 in relation to the location and time component, use of raw survey catch data from multiple  
352 surveys provides a reasonably accurate representation of temporal and spatial variation in  
353 species' abundances (by length category) at large spatial scales. The common dab (Figure  
354 3a/b.), highlights a circumstance in which little variance can be attributed to gear effects, and  
355 we see a consequent small difference in inference in the temporal and spatial trends between  
356 the full and reduced models. The variance explained by the gear is  $<1\%$  while the spatial and  
357 temporal components explain 62.2% of the variance. Thus, this species (by length category)  
358 abundance appears to be less impacted by the effects of gear as the catch rates are likely  
359 driven by the ecology of the fish. The variation that is attributable to gear effects is smaller  
360 than that attributed to space and time in most of our GAMM models, but the nature of the  
361 gear effects are not randomly distributed throughout the study area or throughout the year.  
362 They are instead systematically distributed by seasonal surveys. This regularity in the  
363 differences may impact species distribution inference at large scales. Simulations (S2a) for  
364 species demonstrating substantial movements in distribution attributed 5.7 % of model  
365 variance to gear, even when no gear effect was included. This suggests that some of the  
366 variance associated with location and time may be attributed to gear, but inferences from full  
367 and reduced models were similar. Conversely, when there is a strong gear effect (S2b) then  
368 the full model improves inference of abundance estimates and direction of population centre  
369 of mass movements over the reduced model (Supplemental Material 2).

370

371 Not accounting for gear may lead to incorrect estimates of relative abundance or species'  
372 distributions. Data analysed here suggest that gear effects on catches across disparate surveys  
373 are not uncommon, whereby in half of our full models (127/254), the gear component  
374 explained more than 5% of the total variation in survey catches, while overall variance  
375 explained is >50%. Our examination of the distribution of sole provides demonstration of the  
376 potential importance of controlling for gear effects when attempting to combine data across  
377 surveys for some species. The variance explained by gear in this case was 8.6%, while the  
378 spatial and temporal components of the model accounted for 46.5% of the variance.  
379 Consequently, we found substantial differences in relative abundance trends between models  
380 which control for gear effects compared to reduced models which ignore gear effects in  
381 combining data across surveys (Figures 3d/e, 4, 5). Importantly, failure to control for gear  
382 differences across surveys for this species would mask differences in the spatial distribution  
383 of the stock across commercial fishing areas, as well as mask ecosystem-scale population  
384 shifts to the east (Figure 5). It may be valid to pool across surveys in assessing species  
385 distributions for many species-size combinations; however, there are differences evident  
386 across gear types and it is not clear a priori for which species gear differences matter (Figure  
387 6b). Thus, a sensible workflow when combining data across surveys may be to implement  
388 models that control for gear type as demonstrated here and then subsequently evaluate  
389 whether gear differences account for a substantial portion of the variation in catches.

390

391 Northeast Atlantic waters are currently surveyed by 12 countries carrying out 19 different  
392 surveys designed with individual goals and objectives and using different vessels and a  
393 variety of gears (Table 1). ICES facilitates survey coordination and collaboration through  
394 working groups to make the surveys as comparable as possible. The North Sea bottom trawl  
395 surveys have led the way in terms of minimising gear efficiency issues caused by differences



396 in vessels and by ensuring survey overlap and similarity among gears (ICES, 2015). There is  
397 a large body of work ongoing in ICES survey groups (e.g. WGBEAM, International Bottom  
398 Trawl Survey Working Group; IBTSWG) to minimise survey variability; however, assessing  
399 relative gear efficiency at the scale examined here highlights the need for comparative  
400 experiments to help achieve a more coherent understanding of gear efficiency within fisheries  
401 independent survey data. This is particularly relevant in the Bay of Biscay, where  
402 overlapping or paired tows between the Spanish Baca Trawl and Portuguese Norwegian  
403 Campelen Trawl and the Spanish Baca Trawl and French Grande Overture Vertical Trawl  
404 would help to improve inferences of species relative abundance obtained from these different  
405 gears (Figure 6a). Analyses herein provide further understanding of the differences in gear  
406 efficiency between trawl gears used by different surveys for species sampled across the  
407 northeast Atlantic.

408

409 Information on the abundance and distribution of organisms is a fundamental knowledge  
410 need for fisheries management. Data on predator and prey abundances by different age and  
411 size classes can inform species status assessments as well as provide information on the  
412 interactions among species and size classes, providing understanding about the impact of  
413 fishing on fish communities (Fraser et al., 2007; e.g. Large Fish Indicator). This study  
414 provides an approach to facilitate comparability between catches from different surveys and  
415 gears, providing a framework to assist in integrating data across countries, regions, and  
416 sampling programs towards maximizing the use of available information to inform species'  
417 abundance and spatial distribution assessments.

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424

## 425 **References**

426 Arreguín-Sánchez, F. 1996. Catchability: a key parameter for fish stock assessment. *Reviews*  
427 *in fish biology and fisheries*, 6(2), 221-242.

428 Bivand, R., Keitt, T., and Rowlingson, B. 2019 rgdal: Bindings for the 'Geospatial' Data  
429 Abstraction Library. R package version 1.4-3. [https://CRAN.R-](https://CRAN.R-project.org/package=rgdal)  
430 [project.org/package=rgdal](https://CRAN.R-project.org/package=rgdal)

431 Dann, N., Heessen, H., terHofstede, R. 2005 North Sea Elasmobranchs: distribution,  
432 abundance and biodiversity. Theme Session on Elasmobranch Fisheries Science ICES  
433 CM 2005/N:06

434 EC. 2008. Directive 2008/56/EC of the European Parliament and of the Council of 17 June  
435 2008 establishing a framework for community action in the field of marine  
436 environmental policy (Marine Strategy Framework Directive). *Off. J. Eur. Union* 164,  
437 19–40.

438 EC. 2010. Commission decision of 1 September 2010 on criteria and methodological  
439 standards on good environmental status of marine waters. *Official Journal of the*  
440 *European Union*, L232: 14–24.

441 EC. 2013. Regulation (EU) No 1380/2013 of the European Parliament and of the Council of  
442 11 December 2013 on the Common Fisheries Policy, amending Council Regulations  
443 (EC) No 1954/2003 and (EC) No 1224/2009 and repealing Council Regulations (EC)

444 No 2371/2002 and (EC) No 639/2004 and Council Decision 2004/585/EC. Official  
445 Journal of the European Union, L354: 22–61.

446 EC. 2017. Commission Decision (EU) 2017/848 of 17 May 2017 laying down criteria and  
447 methodological standards on good environmental status of marine waters and  
448 specifications and standardised methods for monitoring and assessment, and repealing  
449 Decision 2010/477/EU. Official Journal of the European Union, 18.5.2017 L 125: 43–  
450 74.

451 Fraser, H. M., Greenstreet, S. P. R., and Piet, G. J. 2007. Taking account of catchability in  
452 groundfish survey trawls: implications for estimating efficiency of survey and  
453 commercial trawl gears estimating demersal fish biomass. ICES Journal of Marine  
454 Science, 64: 1800–1819.

455 Greenstreet, S.P.R., and Moriarty, M., 2017a. Ospar Interim Assessment 2017 Fish Indicator  
456 Data Manual (Relating to for Version 2 of the Groundfish Survey Monitoring and  
457 Assessment Data Product). Scottish Marine and Freshwater Science Report Vol 8 No.  
458 17 DOI: 10.7489/1985-1

459 Greenstreet, S.P.R., and Moriarty, M., 2017b. Manual for Version 3 of the Groundfish Survey  
460 Monitoring and Assessment Data Product. Scottish Marine and Freshwater Science  
461 Report Vol.8 No.18 DOI:10.7489/1986-1

462 ICES. 2015 Manual for the International Bottom Trawl Surveys. Series of ICES Survey  
463 Protocols SISP 10 - IBTS IX. 86 pp.

464 ICES. 2017. Interim Report of the Working Group on Biodiversity Science (WGBIODIV),  
465 6–10 March 2017, Venice, Italy. ICES CM 2017/SSGEPD:01. 14 pp.

466 Kasatkina, S., & Ivanova, V. (2009). Modelling study of catchability properties of research  
467 and commercial trawls to identify sources of uncertainty in resource surveys  
468 indices. *ICES CM*, 1, 13.

469 Kelley, E., & Sherman, K. (2018). Trends of the Large Marine Ecosystem assessment and  
470 management approach as reflected in the literature. *Ocean & Coastal*  
471 *Management*, 155, 104-112.

472 Kotwicki, S., Ressler, P.H., Ianelli, J.N., Punt, A.E., and Horne, J.K., 2018 ‘Combining data  
473 from bottom-trawl and acoustic-trawl surveys to estimate an index of abundance for  
474 semipelagic species’ *Can. J. Fish. Aquat. Sci.* 75: 60–71 (2018)  
475 [dx.doi.org/10.1139/cjfas-2016-0362](https://doi.org/10.1139/cjfas-2016-0362)

476 MacKenzie DI, Nichols JD, Royle JA, Pollock KH, Bailey LL, et al. (2006) Occupancy  
477 estimation and modeling: inferring patterns and dynamics of species occurrence. New  
478 York: Academic Press. 324 p.

479 Maunder, M. N., & Piner, K. R. (2014). Contemporary fisheries stock assessment: many  
480 issues still remain. *ICES Journal of Marine Science*, 72(1), 7-18.

481 Moriarty, M., Greenstreet, S.P.R., Rasmussen, J., de Boois, I., (2019) Assessing the State of  
482 Demersal Fish to Address Formal Ecosystem Based Management Needs: Making  
483 Fisheries Independent Trawl Survey Data ‘Fit for Purpose’ , *Frontiers in Marine*  
484 *Science* 6,162 10.3389/fmars.2019.

485 Moriarty, M., Greenstreet, S.P.R. and Rasmussen, J. (2017). Derivation of Groundfish Survey  
486 Monitoring and Assessment Data Products for the Northeast Atlantic Area. Scottish  
487 Marine and Freshwater Science Report Vol 8. No 16. DOI:10.7489/1984-1

488 Oksanen, J., Blanchet, F. G., Friendly, M., Kindt, R., Legendre, P., McGlenn, D., Minchin,  
489 P.R., O'Hara, R. B., Simpson, G. L., Solymos, P., Henry, M., Stevens, H., Szoecs

490 E., and Wagner, H. (2017). *vegan*: Community Ecology Package. R package version  
491 2.4-5.

492 Ono, K., Ianelli, J. N., McGilliard, C. R., and Punt, A. E. 2017. Integrating data from multiple  
493 surveys and accounting for spatio-temporal correlation to index the abundance of  
494 juvenile Pacific halibut in Alaska. – *ICES Journal of Marine Science*, 75: 572–584.

495 Piet, G. J., van Hal, R., and Greenstreet, S. P. R. 2009. Modelling the direct impact of bottom  
496 trawling on the North Sea fish community to derive estimates of fishing mortality for  
497 non-target fish species. *ICES Journal of Marine Science*, 66: 1985–1998.

498 Poos, J. J., Aarts, G., Vandemaele, S., Willems, W., Bolle, L. J., and van Helmond, A. T. M.  
499 2013. Estimating spatial and temporal variability of juvenile North Sea plaice from  
500 opportunistic data. *Journal of Sea Research*, 75: 118–128.

501 R Core Team (2017). *R: A language and environment for statistical computing*. R Foundation  
502 for Statistical Computing, Vienna, Austria. URL <https://www.R-project.org/>.

503 Somerton, D., Ianelli, J., Walsh, S., Smith, S., Godø, O. R., and Ramm, D. 1999.  
504 Incorporating experimentally derived estimates of survey trawl efficiency into the  
505 stock assessment process: a discussion. *ICES Journal of Marine Science*, 56: 299–  
506 302.

507 Spalding, Mark D., Helen E. Fox, Gerald R. Allen, Nick Davidson, Zach A. Ferdana, M. A.  
508 X. Finlayson, Benjamin S. Halpern et al., "Marine ecoregions of the world: a  
509 bioregionalization of coastal and shelf areas." *BioScience* 57, no. 7 (2007): 573-583.

510 Thorson, J.T., Ianelli, J.N., Larsen, E.A., Ries, L., Scheuerell, M.D., Szuwalski, C., and  
511 Zipkin, E.F., 2016. Joint dynamic species distribution models: a tool for community  
512 ordination and spatiotemporal monitoring. *Global Ecology and Biogeography*, (2016)  
513 25, 1144–1158

514 US. 1996. Magnuson-Stevens Fishery Conservation and Management Sustainable Fisheries  
515 Act of 1996 Public Law 104-297

516 US. 2006. Magnuson-Stevens Fishery Conservation and Management Reauthorization Act of  
517 2006 Public Law 109-479, 109th Congress, Stat. pp. 3575–3665.

518 Walker, N. D., Maxwell, D. L., Le Quesne, W. J. F., and Jennings, S. 2017. Estimating  
519 efficiency of survey and commercial trawl gears from comparisons of catch-ratios. –  
520 ICES Journal of Marine Science, doi:10.1093/icesjms/fsw250.

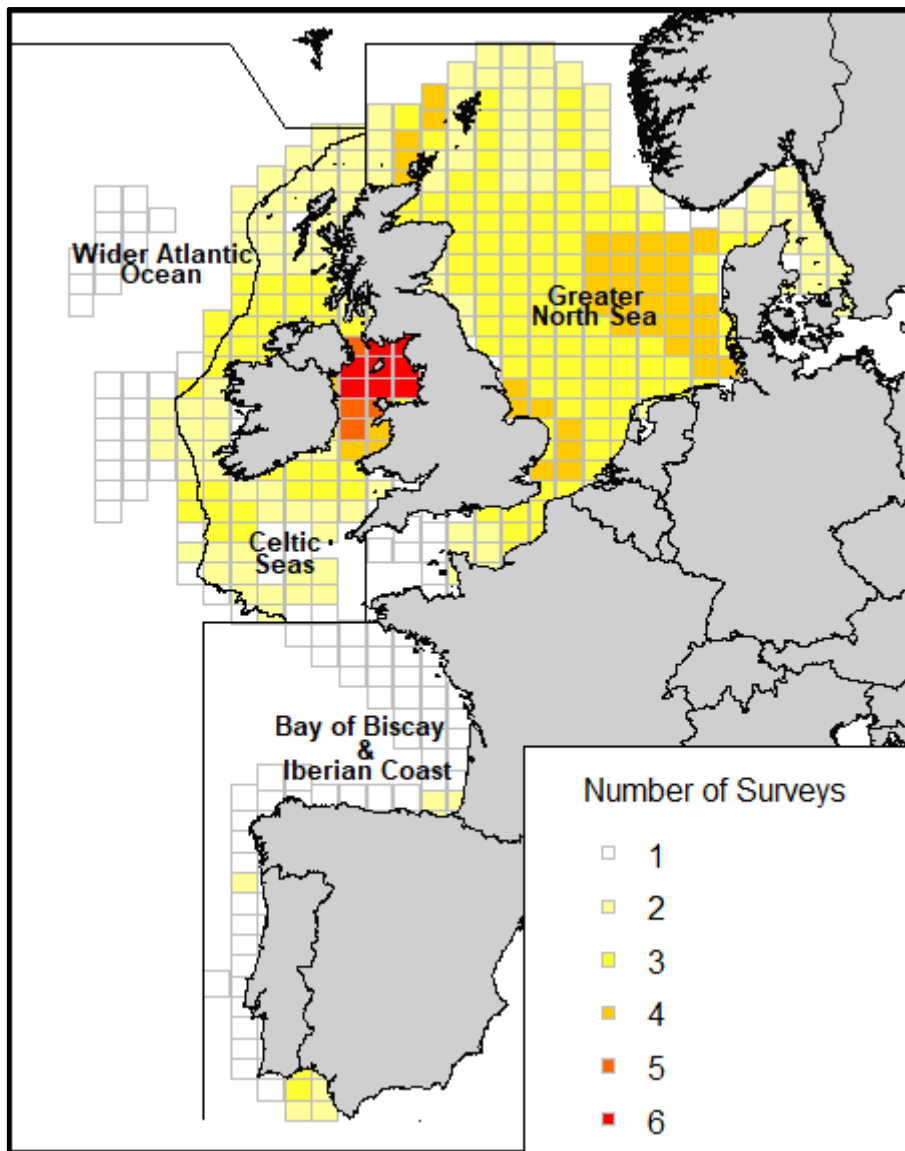
521 Ward, J. H., Jr. (1963), "Hierarchical Grouping to Optimize an Objective Function", *Journal*  
522 *of the American Statistical Association*, 58, 236–244.

523 Wood, S.N. (2004) Stable and efficient multiple smoothing parameter estimation for  
524 generalized additive models. *Journal of the American Statistical Association*. 99:673-  
525 686.

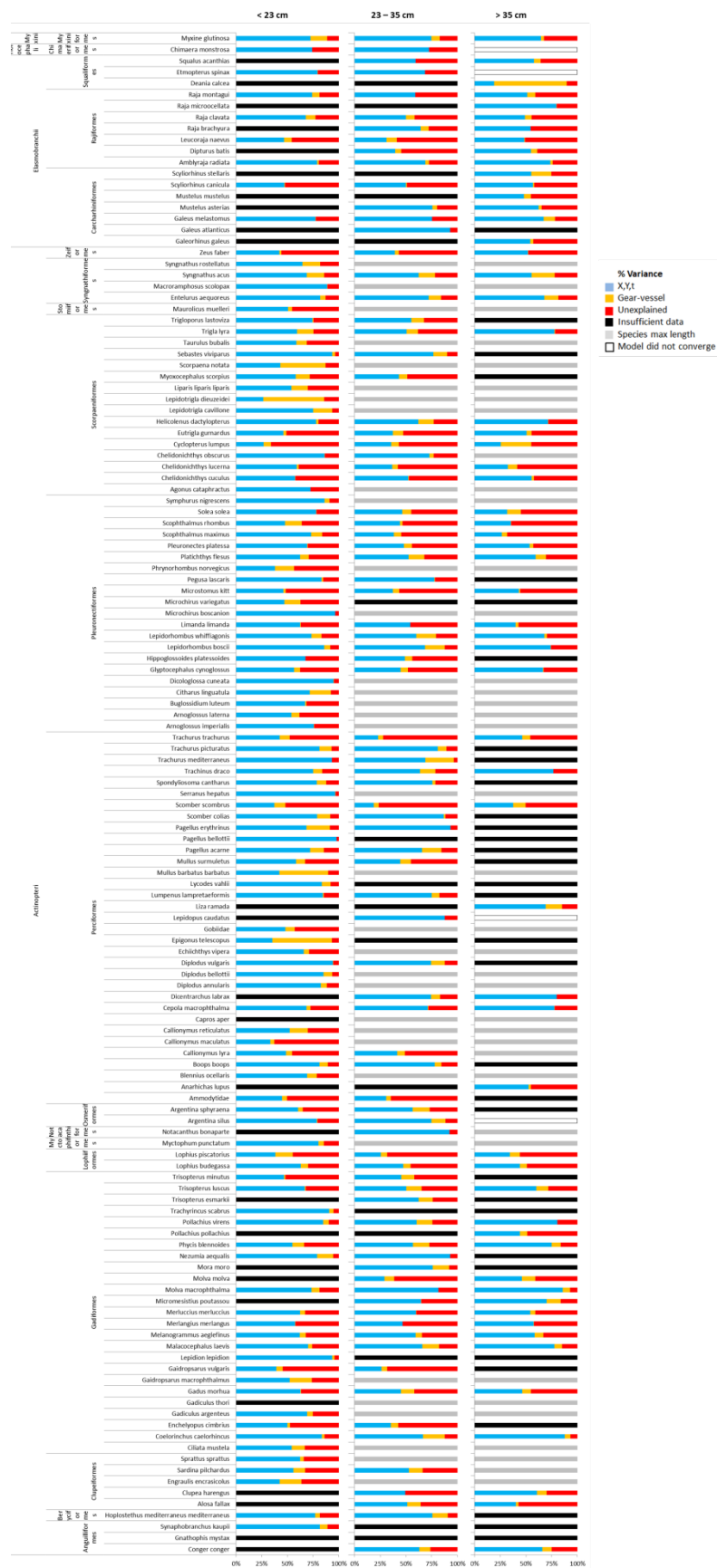
526 Wood, S.N. (2008) Fast stable direct fitting and smoothness selection for generalized additive  
527 models. *Journal of the Royal Statistical Society (B)* 70(3):495-518

528 Wood, S.N. (2011) Fast stable restricted maximum likelihood and marginal likelihood  
529 estimation of semiparametric generalized linear models. *Journal of the Royal*  
530 *Statistical Society (B)* 73(1):3-36

531 Zhou, S., Klaer, N. L., Daley, R. M., Zhu, Z., Fuller, M., and Smith, A. D. M. 2014.  
532 Modelling multiple fishing gear efficiencies and abundance for aggregated  
533 populations using fishery or survey data. *ICES Journal of Marine Science*, 71: 2436–  
534 2447.



535  
 536 **Figure 1:** Fisheries independent survey coverage across the northeast Atlantic. Thick black  
 537 line shows Oslo/Paris convention (OSPAR) boundaries. Number of surveys operating in each  
 538 ICES statistical rectangle is depicted by a different colour. See Table 1 for list of surveys.

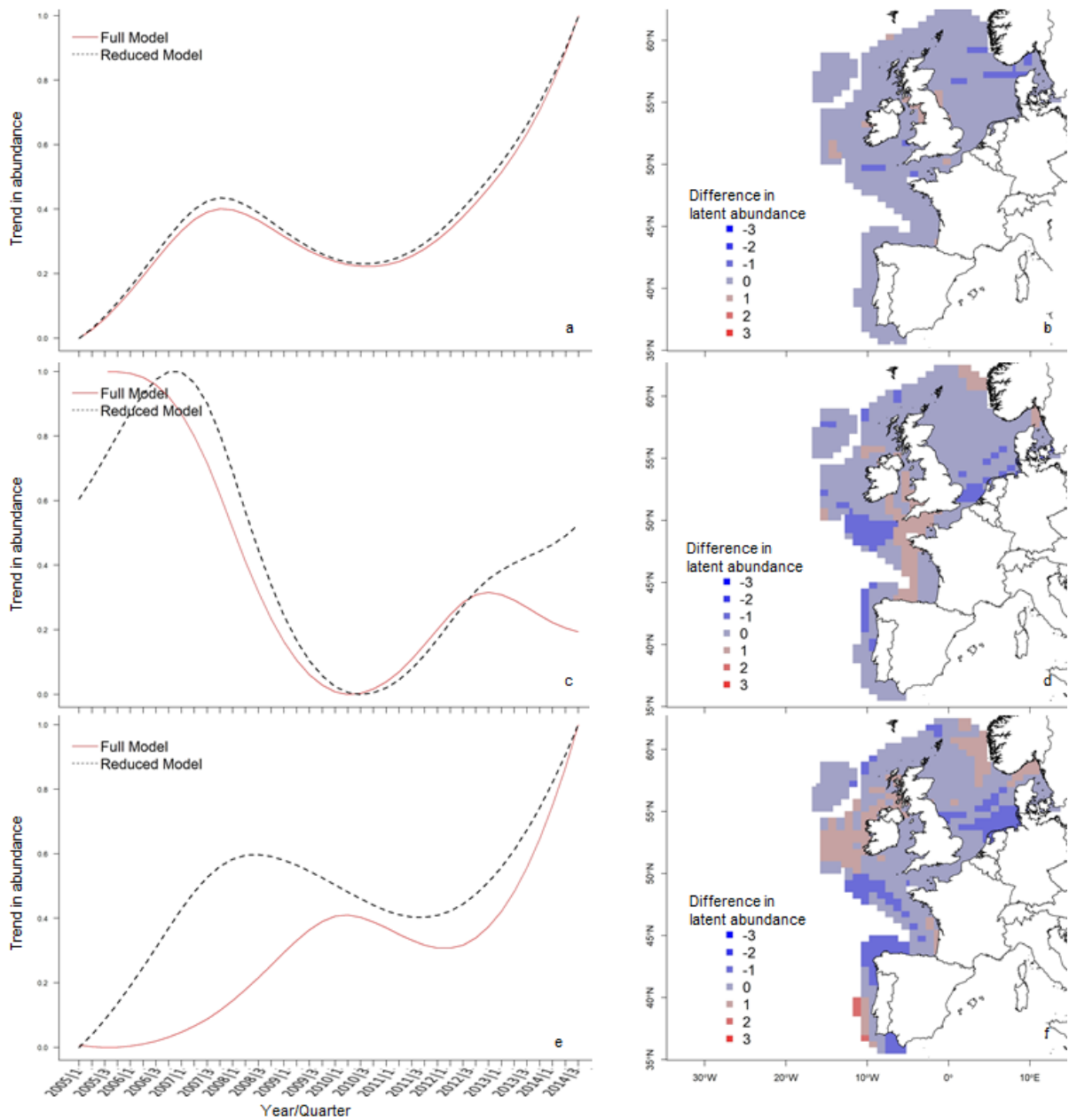


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Figure 2: Summary of the proportion of variance explained from full model components for each length category (<23cm; 23 - 35cm and >35cm) and species, grouped in taxonomic order. X, Y and time (t) variance components are represented by blue bars, gear-vessel

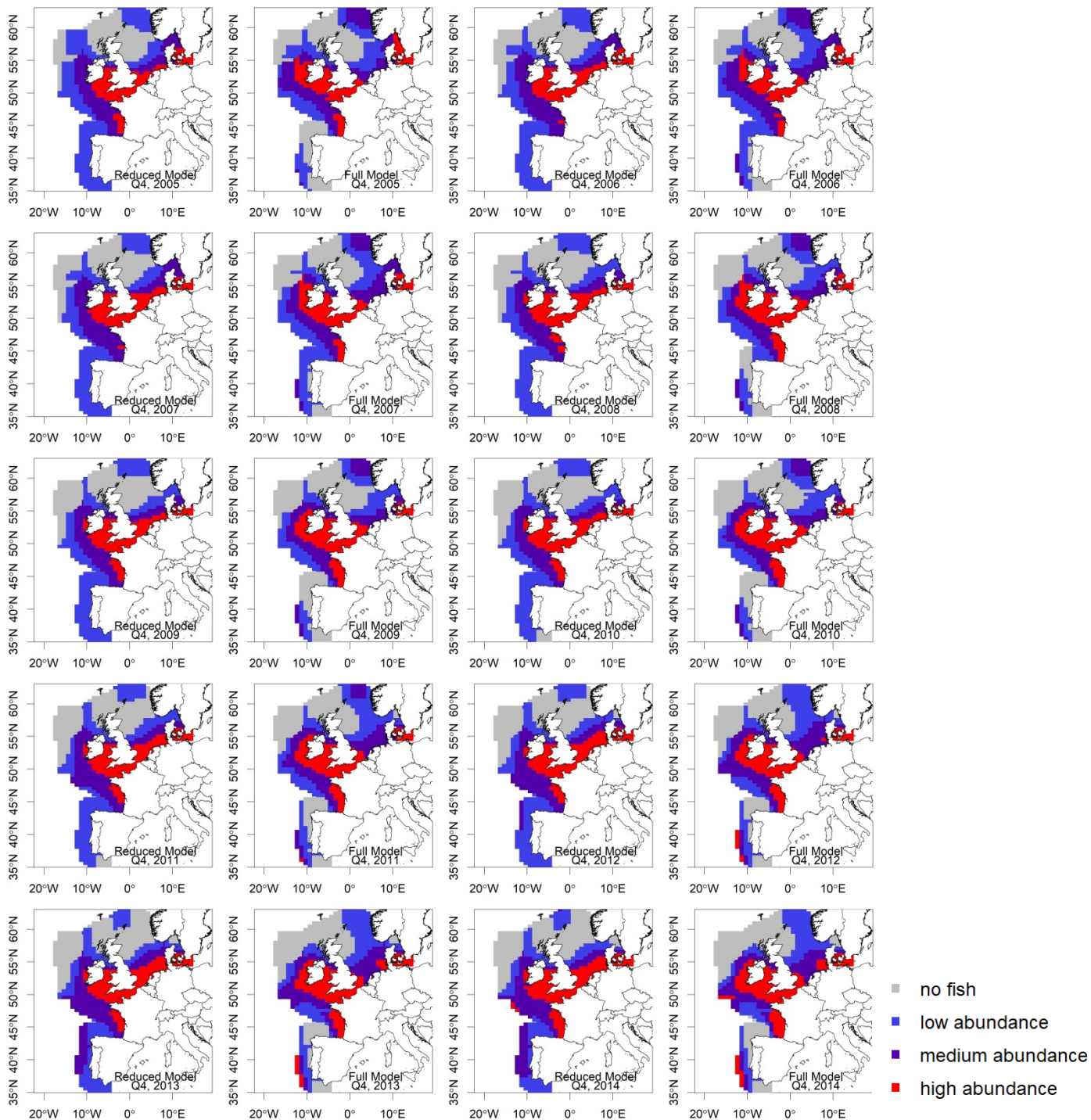


543 components by orange bars, and unexplained variance by red bars. Black bars indicate  
 544 insufficient data to fit a model for a given species-size combination, and white bars indicate  
 545 model convergence failed. Finally, grey bars indicate a given length size bin is larger than  
 546 the maximum observed length of a species.

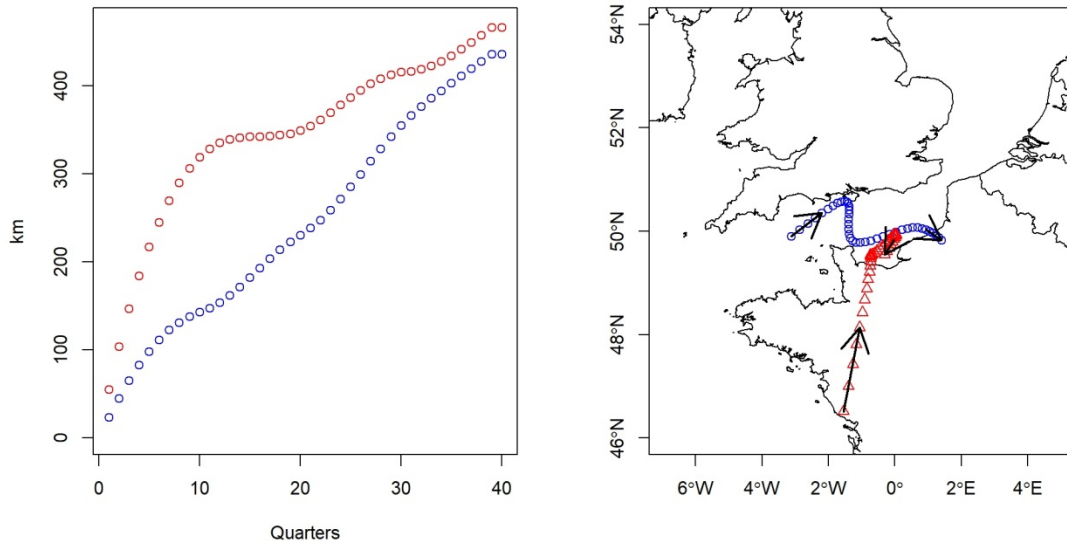


547  
 548 *Figure 3: Top row (a,b): Common dab (*Limanda limanda*) < 23cm, highlighting an example*  
 549 *of a species where the reduced and full model perform similarly as the variance explained by*  
 550 *gear is very small (0.007%). Middle row (c,d) Thorny skate (*Amblyraja radiata*) 23 – 35cm,*  
 551 *highlighting an example of a species with between 1-5% variance explained by gear. Bottom*  
 552 *row (e,f) Sole (*Solea solea*) 23-35cm, highlighting an example of a species with >5%*  
 553 *variance explained by gear. Left column (a,c,e): Estimated domain-wide species' abundance*  
 554 *trends for the full model which controls for gear differences across surveys, versus the*  
 555 *reduced model which does not control for gears. A large discrepancy between the curves*  
 556 *indicates gear differences across surveys may impact inference about species' abundance*  
 557 *and distributions. Right column (b,d,f): Differences in predicted species' relative mean*

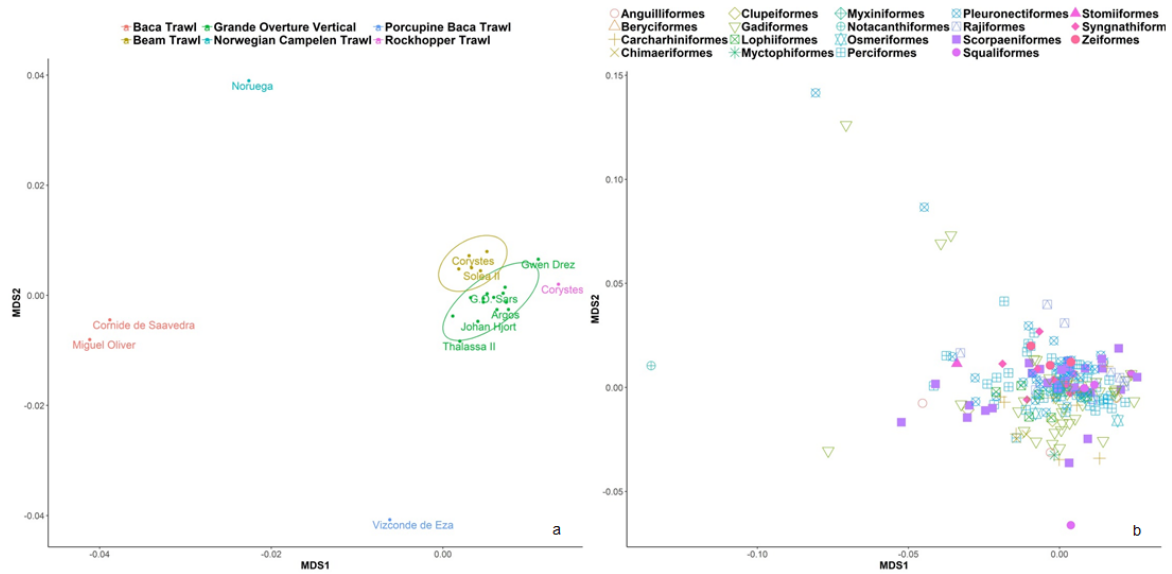
558 abundance between the full and reduced models. Dark colours represent large discrepancies  
 559 between the models, indicating differences in gears across surveys may influence estimated  
 560 species' distributions if not accounted for.  
 561



564 *Figure 4. Spatial-temporal pattern in quarter 4 (Oct.-Dec.) for each year of sole (Solea*  
 565 *solea) 23-35cm from the reduced model on the left and the full model on the right.*  
 566 *Abundance is depicted as “low” in the 1<sup>st</sup>-2<sup>nd</sup> quantile, “medium” in the 2<sup>nd</sup> – 3<sup>rd</sup>*  
 567 *and high in the 3<sup>rd</sup>-4<sup>th</sup> quantile.*



568  
 569 *Figure 5. Summary of difference in inference from the spatial-temporal pattern of sole*  
 570 *(Solea solea) 23-35cm from the full and reduced models. (a.) Cumulative movement from the*  
 571 *centre of mass from the start of the time series for the full model (blue circles) and the*  
 572 *reduced model (red triangles). (b.) Centre of mass of the abundance of the fish from the full*  
 573 *model (blue circles) and the reduced model (red triangles).*  
 574



575  
 576 *Figure 6. Non-metric multidimensional scaling (nMDS) plots describing how the gear-vessel*  
 577 *coefficients varied by survey or by taxonomic grouping (Stress = 0.102). (a) Gear coefficients*  
 578 *grouped by trawl type (colours) and survey research vessel name (labels). Points more*  
 579 *closely situated are more similar in terms of their gear-vessel coefficients. Ellipses indicate*  
 580 *the 95% confidence intervals for clusters of each gear type. (b) Gear coefficients grouped by*  
 581 *taxonomic order as a proxy for species' functional form.*

582

583

Survey Acronym	DATRAS Acronym	Subregion	Country	Start Year	End Year	Quarter	Vessels	Gear Type	Mesh size (mm)	Haul Duration (min) $\bar{x} \pm s$	Distance Towed (km) $\bar{x} \pm s$	Wing Swept Area (km <sup>2</sup> ) $\bar{x} \pm s$	Data Source	DOI
GNSIntOT1	IBTS	Greater North Sea	International	1983	2017	1	Multiple Ships	Otter (GOV)					DATRAS	10.7489/1922-1
							G.O.Sars		20	29±4	3.4±0.6	0.06±0.01		
							Argos		20	30±4	3.6±0.5	0.07±0.01		
							Dana		20	30±2	3.6±0.3	0.07±0.01		
							Dana (Sweden)		20	30±2	3.4±0.3	0.07±0.01		
							CEFAS Endeavour (Netherlands)		20	29±3	3.6±0.4	0.0±0.017		
							Haakon Mosby		20	29±3	2.9±0.3	0.06±0.1		
							Mimer		20	29±3	3.3±0.3	0.06±0.01		
							Scotia III		20	31±6	3.6±0.7	0.07±0.01		
							Thalassa II		20	30±1	3.6±0.4	0.06±0.01		
GNSIntOT3	IBTS	Greater North Sea	International	1998	2016	3	Multiple Ships	Otter (GOV)					DATRAS	10.7489/1923-1
							Argos		20	30±1	3.5±0.2	0.07±0.01		
							Dana		20	29±4	3.6±0.5	0.07±0.01		
							Dana (Sweden)		20	30±1	3.4±0.2	0.07±0.01		
							CEFAS Endeavour		20	29±3	3.5±0.4	0.07±0.01		
							Haakon Mosby		20	27±5	3.2±0.6	0.07±0.01		
							Johan Hjort		20	27±6	3±0.8	0.06±0.02		
							Scotia III		20	29±4	3.3±0.5	0.06±0.01		
							Walther Herwig III		20	29±4	3.6±0.7	0.07±0.01		
							GNSFraOT4	FR CGFS	Greater North Sea	France	1988	2016		
Thalassa II		20	29±2	3.4±0.3	0.05±0.01									
		20	29±3	2.9±0.5	0.03±0.01									
CSScoOT1	SWC-IBTS	Celtic Sea	Scotland	1985	2016	1	Scotia II	Otter (GOV)	20	56±10	7.4±1.7	0.15±0.03	DATRAS	10.7489/1957-1
							Scotia III		20	30±6	3.4±0.7	0.07±0.01		
CSScoOT4	SWC-IBTS	Celtic Sea	Scotland	1997	2016	4	Scotia II	Otter (GOV)	20	56±10	6.6±1.7	0.13±0.03	DATRAS	10.7489/1924-1
							Scotia III		20	29±3	3.4±0.4	0.06±0.01		
CSIreOT4	IE-IGFS	Celtic Sea	Ireland	2003	2016	4	Celtic Explorer	Otter (GOV)	20	30 ± 2	3.6±0.3	0.07±0.01	DATRAS	10.7489/1925-1
CSNIrOT1	NIGFS	Celtic Sea	Northern	1992	2016	1	Corystes, Lough	Otter (ROT)					DATRAS	10.7489/1961-

			Ireland				Foyle								1
							Corystes		20	58±9	5.3±0.9	0.08±0.01			
							Lough Foyle		20	59±6	5.±0.6	0.08±0.01			
CSNIrOT4	NIGFS	Celtic Sea	Northern Ireland	1992	2016	4	Corystes, Lough Foyle	Otter (ROT)					NDB (92-07) DATRAS (08-15)	10.7489/1962-1	
							Corystes		20	19±1	1.9±0.01	0.03±0.02			
							Lough Foyle		20	50±18	4.7±1.6	0.07±0.02			
CS/BBFraOT4	EVHOE	Celtic Sea/Bay of Biscay	France	1997	2016	4	Thalassa II	Otter (GOV)	20	30 ± 1	3.6±0.2	0.07±0.01	NDB (92-07) DATRAS (08-15)	10.7489/1958-1	
BBIC(n)SpaOT4	SP-North	Bay of Biscay and Iberian Coast	Spain	1993	2014	4	F deP Navarro	Otter (BACA)	20	30		0.05	NDB	Not released, no DOI	
							Cornide Saavedra de		20	30	2.7±0.1				
							F deP Navarro	Otter (BACA)	20	60	2.8±0.2				
BBIC(s)SpaOT1	SP-ARSA	Bay of Biscay and Iberian Coast	Spain	1990	2015	1	F deP Navarro	Otter (BACA)	20	60	5.6±0.2	0.1 ± 0.02	NDB	Not released, no DOI	
							Cornide Saavedra de		20	60	5.6±0.4	0.1±0.01			
BBIC(s)SpaOT4	SP-ARSA	Bay of Biscay and Iberian Coast	Spain	1997	2014	4	F deP Navarro	Otter (BACA)	20	60	5.5±0.3	0.09±0.02	NDB	Not released, no DOI	
							Cornide Saavedra de		20	60	5.5±0.3	0.1±0.01			
BBICPorOT4	PT-IBTS	Bay of Biscay and Iberian Coast	Portugal	2001	2014	4	Capricornio, Noruega	Otter (NCT)	20	29±3	3.1±0.4	0.05±0.01	DATRAS	10.7489/1963-1	
WAScoOT3	Rockall	Wider Atlantic	Scotland	1999	2016	3	Scotia III	Otter (GOV)	20				DATRAS	10.7489/1960-1	
										30±3	3.4±0.4	0.07±0.01			
WASpaOT3	SP-PORC	Wider Atlantic	Spain	2001	2015	3	Vizconda de Eza	Otter (PBACA)	20	24 ± 4	2.7±0.5	0.07±0.02	NDB	Not released, no DOI	
GNSNetBT3	BTS	Greater North Sea	The Netherlands	1999	2016	3	Isis, Tridens II	Beam (8m)					DATRAS	10.7489/1967-1	
							Isis		40	30±2	3.8±0.3	0.03			
							Tridens II		40	34±11	4.5±1.4	0.04±0.01			

GNSEngBT3	BTS	Greater North Sea	England	1990	2016	3	Corystes	Beam (4m)	40	29±3	3.7±0.6	0.01	DATRAS	10.7489/1966-1
							Endevour		40	28±4	3.5±0.6	0.01		
GNSGerBT3	BTS	Greater North Sea	Germany	1998	2016	3	Solea I	Beam (7m)	40	30±3	3.5±0.5	0.03	DATRAS	10.7489/1965-1
							Solea II		40	30 ± 2	3.3±0.3	0.02		
CSEngBT3	BTS VIIa	Celtic Sea	England	1993	2015	3	Corystes, Endevour	Beam (4m)					DATRAS	10.7489/1964-1
							Corystes		40	28 ± 5	3.8±0.5	0.02		
							Endevour		40	28 ± 4	3.5±0.6	0.01		

584 *Table 1. List of individual surveys considered in the derivation of the Oslo/Paris convention (OSPAR) Groundfish Survey Monitoring and Assessment data*  
585 *products. Survey acronyms reflect sub-region/country/gear/quarter, except CS/BB in the French EVHOE survey acronym to denote a survey that extends*  
586 *across two sub-regions, the Celtic Seas and Bay of Biscay. Data product start and end years reflect the period when surveys were deemed sufficiently*  
587 *established with consistent standardised methodology (Moriarty et al., 2017). NDB refers to national database. For this study we subset the data from 2004 –*  
588 *2015 for continuous spatial coverage across the northeast Atlantic, the information on mesh size, haul duration, distance towed, and wing swept area reflect*  
589 *the data included in this study from 2004-2015 where  $\bar{x}$  is the sample mean and  $s$  is the sample standard deviation.*

590

#### 591 **Notes on fishing gear exceptions**

592 S = Standard Gear      B = Bobbins used      D = Double Sweeps      I2 = Ground gear D with 16-inch bobbins      R = Rockhopper

#### 593 Grande Overture Vertical Trawl

- 594 1. Scotland uses R.V. Scotia III on five surveys WAScoOT3; CSScoOT4; CSScoOT1; GNSIntOT3; GNSIntOT1. For the west coast surveys (CSScoOT4 /  
595 CSScoOT1/ WAScoOT3) they use a “S” and “I2” gear for to deal with rocky habitat. In North Sea surveys (GNSIntOT3; GNSIntOT1), Scotland uses  
596 an “S” and a “B” exception.
- 597 2. Sweden uses a standard GOV (“S”) on R.V Argos and R.V. Mimer in both North Sea surveys (GNSIntOT3 and GNSIntOT1).
- 598 3. Denmark uses an “S” gear and an “R” exception in both surveys on R.V. Dana II in both North Sea surveys (GNSIntOT3 and GNSIntOT1).
- 599 4. England uses a standard GOV (“S”) gear in the North Sea (GNSIntOT3) on R.V. CEFAS Endeavour.
- 600 5. The Netherlands uses a standard GOV (“S”) gear in the North Sea (GNSIntOT1) on R.V. Tridens II. R.V. CEFAS Endeavour was used in quarter 1 by  
601 Netherlands when Tridens broke down.
- 602 6. Norway uses an “S” gear and “D” exception on R.V. G.O. Sars and R.V. Johan Hjort in the North Sea (GNSIntOT3/ GNSIntOT1). When the R.V  
603 Haakon Mosby has been used only a standard gear is noted.
- 604 7. France uses a GOV gear in the North Sea (GNSFraOT4) on R.V. Gwen Drez, no exception is noted, however, the gear is smaller than the standard  
605 gear in the North Sea. France uses Thalassa II on two surveys CS/BBFraOT4 and GNSIntOT1. For the west coast surveys (CS/BBFraOT4) they use  
606 ground gear “D” while in the North Sea surveys (GNSIntOT1), a standard gear is used.
- 607 8. Germany uses a standard gear on R.V Walther Herwig III in the North Sea (GNSIntOT3/ GNSIntOT1).
- 608 9. Ireland uses an “S” and “I2” gear for west coast survey (CSIREOT4) to deal with rocky habitat in line with Scotland on R.V. Celtic Explorer.

#### 609 Beam Trawl

- 610 10. The Netherlands uses R.V. *Tridens II* and R.V. *Isis* in the GNSNetBT3 survey. Both ships use an 8m beam with a tickler but *Tridens II* has a different set  
611 up to *Isis*.
- 612 11. Germany uses a 7m Beam trawl with a 5m tickler chain on R.V. *Solea II* during GNSGerBT3.
- 613 12. England uses a 4m Beam trawl during both her CSEngBT3 and GNSEngBT3 surveys on R.V. *Corystes* and CEFAS *Endeavour* in 2014 and 2015 with  
614 the same rigging on both ships.
- 615 Rockhopper Trawl
- 616 13. The Rockhopper Otter Trawl is used by Northern Ireland in the CSNIrOT4 / CSNIrOT1 on R.V. *Corystes*.
- 617 Baka Trawl
- 618 14. Spain uses a Baka trawl on 3 surveys (BBIC(s)SpaOT4 / BBIC(s)SpaOT1 / BBIC(n)SpaOT4) on R.V. *Cornide de Saavedra*.
- 619 15. Spain uses a Porcupine Baka trawl on 1 survey (WASpaOT3) on R.V. *Vizconde de Eza*.
- 620 Norwegian Campelen Trawl
- 621 16. Portugal reports B and R gear exceptions on R.V. *Noruega*.

622 **Supplemental Material 1:**

623 **Table S1.1** Variance explained (%) and AIC scores for all models in all length classes. This  
624 includes a folder containing the same information in Figure3 for every species/length  
625 combination.

626 **Supplemental Material 2:** Can GAMMs differentiate effects of gear efficiency from spatial  
627 and temporal variation in abundance in demersal fish?

628 **Supplemental Material 3:**

629 **File S3.1:** Example of R Scripts for fitting model

630 **File S3.2:** Example of R Scripts for generating data and fitting simulated models (S1  
631 conditions only)