

1 Unraveling conflicting density- and distance-dependent effects on plant

2 reproduction using a spatially-explicit approach

- 3 José M. Fedriani^{1,2,3}, Thorsten Wiegand¹, Gemma Calvo², Alberto Suárez-Esteban², Miguel
- 4 Jácome², Magdalena Żywiec⁴, and Miguel Delibes²
- ⁵ ¹Department of Ecological Modelling, Helmholtz Centre for Environmental Research GmbH -
- 6 UFZ. Permoserstrasse 15, 04318 Leipzig, Germany.
- ⁷ ²Department of Conservation Biology, Estación Biológica de Doñana (EBD C.S.I.C.),
- 8 c/Americo Vespucio s/n, 41092 Seville, Spain.
- ³Technical University of Lisbon, Institute of Agronomy, Centre for Applied Ecology, Tapada
- 10 da Ajuda, 1349-017 Lisboa, Portugal.
- ⁴Polish Academy of Sciences, Institute of Botany, ul. Lubicz 46, 31 512 Krakow, Poland
- 12 **Correspondence author.** E-mail:fedriani@ebd.csic.es
- 13 **Running title:** Conflicting density- and distance-dependent effects

15 Summary

Density- and distance-dependent (DDD) mechanisms are important determinants of plant
 reproductive success (PRS). Different components of sequential PRS can operate either in the
 same or in different directions and thus reinforce or neutralize each other, and they may also
 operate at different spatial scales. Thus, spatially-explicit approaches are needed to detect
 such complex DDD effects across multiple PRS components and spatial scales.

2. To reveal DDD effects of different components of early PRS of the Iberian pear (*Pyrus bourgaeana*) sampled over three consecutive years, we used marked point pattern analysis.
Our special interest is to identify conflicting processes that regulate populations at different spatial scales, e.g. whether DDD on fruit initiation and on fruit development acted in opposite directions. To evaluate the significance of observed mark correlation functions based on empirical data (e.g. fruiting success) we compared them to expectations given by spatially-explicit null models.

3. Diverse DDD processes affected several aspects of PRS in a variable extent over the three 28 29 seasons. First, early fruit set was higher for individuals with more neighbors at small 30 distances (i.e. up to 40m). However, late P. bourgaeana fruit set decreased with increasing 31 number of nearby neighbors, but these effects canceled for overall fruit set that did not show DDD effects. Second, the absolute number of fruits produced (crop sizes) by trees showed 32 positive density dependence in 2011 and 2012 but not in 2013. Finally, the total number of 33 seeds produced did not show DDD effects, indicating that conflicting demographic processes 34 can disrupt the initial spatial pattern of tree investment in reproduction. 35

4. Synthesis: Understanding complex spatial effects of density- and distance-dependent
(DDD) processes requires dissection of component processes to attain the complete picture
since contrasting DDD processes may be hidden behind a single cumulative measure of
reproductive success. The combination of novel and classic mark correlation functions used
here constitute a powerful spatially-explicit tool that can be broadly applied to unravel
conflicting mechanisms of DDD regulating the persistence of sessile organisms at a range of

spatial scales. Our findings help to explain why some authors failed to find expected DDD of
PRS and highlight the importance of detailed multi-year field studies on plant reproductive
success.

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- 46 Key words: aggregation, conflicting effects, density dependence, Doñana National Park,
- 47 mark correlation functions, plant-animal interactions, population regulation, pollination, plant

48 population and community dynamics, small populations

50 INTRODUCTION

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Long-term persistence of natural populations is critical for the maintenance of community 52 structure, ecosystems functioning, and essential ecosystem services. Both plant and animal 53 populations are frequently regulated by density and distance dependence of their demographic 54 55 parameters that impose bounds on population size (Strong 1986, Hixon et al. 2002). Contrasting density- and distance-dependent processes can operate during the often intricate 56 life cycles of organisms sequentially or even simultaneously (Roughgarden et al. 1988). 57 Diverse proximate mechanisms of density and distance dependence can thus act during 58 different ontogenic stages, spatial scales, and contrasting directions (e.g. positive vs. negative 59 density-dependence; see Courchamp et al. 2008 for a review). 60 61 In plant populations, the spatial distribution of individuals determines how a species uses 62 63 resources, how it is used as a resource, and how it reproduces (Condit et al. 2000, Wiegand et al. 2007, Fedriani et al. 2010). For instance, isolated individuals are less often visited by 64 effective pollinators and thus produce fewer fruits than more aggregated individuals (i.e. 65 positive density-dependence; Nielsen & Ims 2000, Aizen & Vázquez 2006, Fox 2007). Such 66 negative effects of isolation on fruit initiation can be amplified or attenuated by density 67 68 dependence in other demographic processes acting on subsequent plant life stages (e.g., fruit development). In particular, high conspecific density may either facilitate encounter and 69 exploitation by seed predators and pathogens (Janzen 1970) or promote predator satiation and 70 71 thus attenuate the *per capitat* chances of predation (Augspurger 1981). 72

Competition for resources or pollinators can also lower fitness in high-density neighborhoods
(Harper 1977, Spigler & Chang 2008, Gunton & Kunin 2009). Recent investigations in selfincompatible species have shown that negative density- and/or distance-dependence is most

likely when mating tends to occur among highly related neighbors, which results in low 76 77 quality of available pollen and lessened plant reproductive success (hereafter, PRS) (Ishihama et al. 2006, Aizen & Harder 2007). Whether plant aggregation has an overall net positive or 78 79 negative effect on individual fitness or whether partial effects cancel each other is an important question necessary to fully understand how plant populations are regulated. The 80 answer requires the identification of potentially conflicting density- and distance-dependent 81 82 (hereafter, DDD) effects that act upon different life stages, and may also act over different spatial scales within a population. Like other ecological and evolutionary processes (Levin 83 1992, Peterson et al. 1998, Thompson 2013), density and distance dependence of PRS vary 84 85 with spatial scale and the distance between neighboring conspecifics. Surprisingly, however, few studies have investigated such processes in a spatially-explicit way across spatial scales 86 and on multiple components of plant reproduction (Knight et al. 2005, Gunton & Kunin 2009, 87 88 Gómez et al. 2009). Spatially-explicit approaches are needed to reveal the complexity of density and distance dependent effects that operate simultaneously over multiple spatial scales 89 90 and reproductive success components.

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The analysis of marked spatial point patterns (Illian et al. 2008, Wiegand & Moloney 2014) 92 93 constitutes a powerful, albeit rarely used approach in ecology that allows identification of distance- and density-dependence of PRS over a range of spatial scales. This technique 94 involves statistical analysis of mapped point patterns (e.g., reproductive plants) and additional 95 proprieties of the plants, called "marks" that represent, for example, PRS components (e.g., 96 fruit initiation and development successes, number of seeds per fruit, etc). Analysis of such 97 quantitatively marked patterns allows detection of correlations in PRS across a range of 98 spatial scales (e.g., Illian et al. 2008, Getzin et al. 2011, Raventós et al. 2011, Wiegand et al. 99 2013, Fedriani & Wiegand 2014). This allows us to find out if, for example, whether 100 pollination success is larger for trees that have more neighbors compared to trees with fewer 101

102 neighbors. For detecting such density-dependent effects we specifically tailored a function 103 that directly measures the correlation between plant density and PRS. This new "density 104 correlation function" $D_{mK}(r)$ estimates the correlation coefficient between the reproductive 105 success of all trees and the density of neighbors located within distance *r* of a specific tree. 106

107 In this study, we investigate the occurrence of possibly conflicting DDD processes in different 108 components of pollination and fruiting success of the self-incompatible Iberian pear Pyrus 109 bourgaeana in a Mediterranean area of southern Spain. Our data comprise three consecutive reproductive seasons. The low-density P. bourgaeana population studied here experiences 110 111 severe seed dispersal limitation that leads to small-scale aggregated establishment beneath reproductive trees and eventually to adult tree clustering (Fedriani et al. 2010). These tree 112 clusters are readily located and exploited by diverse pollinators (Żywiec et al. 2012). 113 114 Therefore, we hypothesized that aggregated trees would experience higher pollination success (i.e. fruit initiation) than isolated trees which are more difficult to locate by pollinators. Also, 115 because most recruitment occurs near parent trees and high-density neighborhoods tend to be 116 117 composed by highly related individuals (Rodríguez 2009), we predicted more fruit abortion and lower fruit development (i.e. fruits that complete development) for trees with close 118 119 neighbors than for isolated trees (Herlihy & Eckert 2004, Aizen & Harder 2007). More specifically, we seek to answer the following three questions: 1) Do measures of relative 120 reproductive success (percentage of flowers that set and develop fruit) show DDD effects and, 121 if so, at what spatial scale? 2) What is the net effect of density and distance among neighbors 122 on absolute measures of reproductive success of P. bourgaeana (i.e., crop size, total number 123 of seeds per tree)? and 3) Are the DDD effects of different reproductive components 124 125 consistent in sign and spatial scale, and among reproductive seasons? To answer these questions, and to progress our ability to unravel density from distance dependent mechanisms 126 acting on PRS, we use standard mark correlation statistics in combination with a new density 127

correlation function that quantifies the correlation between a mark of the plants (e.g. a
measurement of PRS) and the number of conspecific neighbors within a given distance *r*.

131 MATERIALS AND METHODS

132 *Study system*

Pyrus bourgaeana (Rosaceae) is a small tree distributed across the southern Iberian Peninsula 133 and northern Morocco (Aldasoro et al. 1996). Our focal population is located in Doñana 134 National Park (510 km²; 37°89′N, 6°26′W), on the west bank of the Guadalquivir River 135 estuary in southwestern Spain. The climate is Mediterranean sub-humid, characterized by 136 dry, hot summers (June-September) and mild, wet winters (November-February). In the 137 Doñana area, this tree population occurs at low densities (< 1 ind/ha) across the landscape but 138 tree local distribution varies from highly isolated to strongly aggregated (6-8 reproductive 139 140 individuals within ~25 m; Fedriani et al. 2010). The establishment and persistence of such tree clusters arises from dispersal limitation and spatial contagion of dispersed seeds (Fedriani 141 142 & Wiegand 2014). These processes create considerable relatedness among individuals within clusters (Rodriguez 2011). The fieldwork (see below) was undertaken in an irregular 143 quadrilateral plot of 49-ha in which the locations of all P. bourgaeana reproductive trees were 144 145 known.

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Pyrus bourgaeana is a self-incompatible species that flowers during February-March and is pollinated by a diverse insect assemblage (Żywiec et al. 2012). Flowers are hermaphrodite, showing radial symmetry and usually five oval petals. Fruit crop size strongly varies among individuals and seasons, though each individual usually produces 100-500 fruits that ripen during fall. Mature fruits are globose pomes, comprising 1-5 viable seeds. Pre-dispersal seed predation by invertebrates is generally low (<5%; Fedriani & Delibes 2009). However,</p>

during the long period of fruit ripening (March-September) a considerable fraction of the

154 initial fruit set fails before completing development (Authors *unpublished data*).

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156 *Fruit set, crop size, and total seed number*

Since *P. bourgaeana* trees produce many inflorescences and reproductive success often varies 157 within trees (Herrera 2009), among each pollination season (i.e. 2011, 2012, and 2013) and 158 for each reproductive individual, we monitored several inflorescences to ensure that we attain 159 a good representation of its reproductive success. Specifically, we selected 2-6 branches of 160 each flowering *P. bourgaeana* tree (n = 71, 67, and 67, respectively) within the study plot 161 (most of the 73 trees flowered during all three seasons). On each branch, we tagged from 1-4 162 inflorescences, recorded the number of flowers per inflorescence, and monitored them once 163 per month until complete fruit fall (i.e. early December). During 2011, our monitoring started 164 165 in early May (i.e. after anthesis) and thus we marked 142 infrutescences with 440 recently set fruits. Thus, for this season, we estimated fruit development (i.e. number of fruits that 166 completed their development until early September / number of set fruits). During 2012 and 167 2013 monitoring started early in March before anthesis; we monitored the fate of 5012 and 168 5048 flower buds placed in 529 and 531 inflorescences, respectively. For both seasons, we 169 estimated both fruit initiation (i.e. number of flowers that set fruit / initial number of flowers) 170 as well as fruit development (as defined above). Overall fruit set (fully developed fruits per 171 flower) for each of these seasons was then estimated as the product of fruit initiation and fruit 172 development successes. Though overall fruit set best describes female performance of 173 individual trees, distinguishing between fruit initiation and development at a range of spatial 174 scales can help to identify density dependent mechanisms acting on individual fitness 175 components. 176

During the three seasons, just before the initiation of fruit fall in early September, we estimated the crop sizes of all reproductive trees within the study plot by counting their fruits with the aid of binoculars, when needed. In *P. bourgaeana*, fruits are arranged in discrete clusters which facilitated estimation of reliable crop size by visually counting fruits from the ground (e.g., Janzen et al. 2008). Finally, during 2011, we estimated the total number of seeds produced by each focal tree by quantifying average seed number per fruit (three fruits per tree) and multiplying it by the corresponding crop size (e.g. Spligler & Chang 2008).

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186 Mark-correlation functions

187 To detect DDD effects on different components of *P. bourgaeana* PRS, we used techniques of marked point pattern analysis (Illian et al. 2008; Wiegand & Moloney 2014). One data set 188 comprises for each reproductive P. bourgaeana tree i the coordinates \mathbf{x}_i and given measure m_i 189 190 of reproductive success that can be in our case percentage of flowers that set and develop fruit, crop size, or the number of seeds per tree. This data structure is a univariate 191 quantitatively marked point pattern (Illian et al. 2008, Wiegand & Moloney 2014) where the 192 193 coordinates \mathbf{x}_i represent the univariate point pattern and the measure of reproductive success $m_{\rm i}$ represents a quantitative mark. 194

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Mark correlation functions are based on all (ordered) pairs of trees which have inter-point 196 distances within the small interval (r - h, r + h). The parameter h is called bandwidth and must 197 be wide enough to yield a sufficient number of pairs in each distance class r but small enough 198 to reveal relevant biological detail (Illian et al. 2008). The basic idea of mark correlation 199 200 functions is then to estimate the mean value $c_t(r)$ of a test function $t(m_i, m_i)$ of the two marks m_i and m_j , taken over all (ordered) pairs i-j of trees which have inter-point distances of $r \pm h$. 201 This procedure is then repeated for a range of distances r taken at distance steps Δr to obtain a 202 non-normalized mark correlation function $c_t(r)$ (Illian et al. 2008). To obtain the final mark 203

204 correlation function, the $c_t(r)$ is normalized with the expectation c_t of the test function taken 205 over all pairs of trees, irrespective on their spatial separation:

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$$k_t(r) = c_t(r)/c_t.$$
 (1)

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Many different test functions $t(m_i, m_j)$ are possible (Illian et al., 2008). We use here the r-mark correlation function $k_{m_i}(r)$ which is based on the test function

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$$t(m_i, m_j) = m_i.$$
 (2)

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An estimator of the corresponding non-normalized mark correlation function is given by

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$$\hat{c}_{t}(r) = \hat{\mu}(r) = \frac{\sum_{i=1}^{n} \sum_{j=1,\neq}^{n} m_{i} \times k^{h}(\|\mathbf{x}_{i} - \mathbf{x}_{j}\| - r)}{\sum_{i=1}^{n} \sum_{j=1,\neq}^{n} k(\|\mathbf{x}_{i} - \mathbf{x}_{j}\| - r)}$$
(3)

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where the "box kernel" function $k^h(d)$ yields a value of 1/2h if the two trees with coordinates **x**_i and **x**_j have inter-point distances of $r \pm h$ and zero otherwise (Illian et al. 2008; Wiegand & Moloney 2014). Thus, the denominator of equation (3) yields the number of ordered pairs of trees i and j which are distances of $r \pm h$ apart and therefore, equation 3 yields the mean value of the mark m_i of the first tree i of these pairs. The normalization constant c_t of the r-mark correlation function which is taken over all pairs of points, regardless of their spatial separation, is obtained by replacing in equation 3 $k^h(d)$ by 1/2h:

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$$\hat{c}_{t} = \frac{\sum_{i=1}^{n} \sum_{j=1,\neq}^{n} m_{i}/2h}{\sum_{i=1}^{n} \sum_{j=1,\neq}^{n} (1/2h)} = \frac{1}{n} \sum_{i}^{n} m_{i} = \mu$$
(4)

and yields μ , the mean value of m_i taken over all trees i (Illian et al. 2008).

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Thus, $k_{m.}(r) > 1$ indicates that the PRS of trees that have neighbors at distance *r* is on average larger than the average PRS, indicating a positive effect of nearby *P. bourgaeana* trees on PRS. Conversely, $k_{m.}(r) < 1$ indicates that the PRS component of trees that have neighbors at distance *r* is smaller than the average PRS, indicating a negative effect of nearby trees on PRS (for further methodological details see Illian et al. 2008 and Wiegand & Moloney 2014).

We also used a mark correlation function that characterizes the spatial covariance in PRS of
two *P. bourgaeana* trees separated by distance *r*. The appropriate test function for this
purpose was proposed by Schlather *et al.* (2004):

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$$t(r, m_i, m_j) = [m_i - \mu(r)][(m_j - \mu(r)]]$$
 (5)

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which results in a Moran's *I* like summary statistic $I_{mm}(r)$, a spatial variant of the classical Pearson correlation coefficient (Shimatani 2002). $I_{mm}(r)$ is normalized by the mark variance σ^2 . Thus, the $I_{mm}(r)$ is the straight forward Pearson correlation coefficient between the two variables m_i and m_j defined by the ordered i-j pairs of trees separated by distance $r \pm h$. Note that a test function that adjusts for the mean $\mu(r)$ that considers only pairs of trees separated by at a given distance r, not the population mean μ , is required to yield a summary statistic with the interpretation of a correlation coefficient (Schlather et al. 2004).

Finally, we developed a new function that directly relates the PRS component of a *P*. *bourgaeana* tree to the density of its conspecific neighbors. This "density correlation function" $C_{m,K}(r)$ estimates the classical Pearson correlation coefficient between the reproductive success m_i of a tree and the number of neighbors within distance $r [=\lambda K_i(r)]$. Thus, the density correlation function is based on the following test function:

$$t(r, m_{\rm i}, K_{\rm i}) = [m_{\rm i} - \mu][(\lambda K_{\rm i}(r) - \lambda K(r)], \qquad (6)$$

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where m_i is the PRS of the focal tree i, μ is the mean PRS of the population, λ the overall density of trees in the study area, $\lambda K_i(r)$ the number of neighbors around the focal tree i within distance r, and $\lambda K(r)$ the mean number of neighbors within distance r for all trees. Because correlation functions are invariant under multiplication with a constant it does not matter if we use absolute number of neighbors (e.g., $\lambda K(r)$) or density of neighbors ($\lambda K(r)/\pi r^2$).

The $K_i(r)$ is basically a "local" K-function and the average over all $K_i(r)$'s yields the well 264 known K-function K(r). Estimation of $K_i(r)$ requires an edge correction factor w_i if the 265 neighborhood around tree i is not fully within the observation window, which yields $\pi r^2/A_i$ 266 where A_i is the area of the disk around point i within the observation window. The density 267 correlation function $C_{m,K}(r)$ is normalized by the product of the standard deviations $\sigma_m \sigma_K$ of 268 the marks m_i and the individual K-functions $K_i(r)$, respectively. "C" stands for correlation, 269 "m" for the first mark m_i , and "K" for the second mark $K_i(r)$. In Appendix S1 we use artificial 270 data sets to test the ability of the density correlation function to reveal different known DDD 271 effects (see Fig. S1-S2). 272

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To test whether the data show spatial correlations in the marks, we contrast the three markcorrelation functions of the data to that of repeated simulations of a null model without spatial

correlation in the marks. This null model was implemented by randomly shuffling the marks over all *P. bourgaeana* trees (Illian et al. 2008; Wiegand et al. 2013). We conducted 199 randomizations to estimate simulation envelopes for our summary statistics $k_{m.}(r)$, $I_{mm}(r)$, and $C_{m,K}(r)$, being the 5th lowest and highest values of the summary statistics among the 199 simulations of the null model (e.g., Wiegand et al. 2013, Fedriani & Wiegand 2014). We observe a departure from the null model at particular distances *r* if the summary statistic of the observed data is outside the simulation envelopes.

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To test the overall fit of the random marking null model over a range of distances we used the "Diggle-Cressie-Loosmore-Ford" Goodness-of-Fit (GoF) (Baddeley et al. 2014) for the entire distance interval up to 50m. The GoF test returns a *P*-value that indicates significant departures of the observed mark correlation function from the random marking null model over the distance interval of interest. If a significant departure occurs, we can inspect the plot of the mark correlation function together with the simulation envelopes to identify the specific distances *r* where departures occurred.

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We estimated the two mark correlation functions and the pair correlation function every five metres up to distances of 50m, where the distance bins centered on r = 2.5m, 7.5m, 12.5m, ... included all pairs of trees within distance (r - h, r + h), being h = 2.5m the bandwidth. The distance bins for the cumulative density correlation function yield r = 5m, 10m, 15m, ... All point pattern analyses were conducted with the software *Programita* (Wiegand & Moloney 2014) available at http://programita.org/

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299 RESULTS

300 Aggregation of P. bourgaeana trees

We first estimated the degree of clustering of our focal tree population. The non-cumulative 301 neighborhood density varied with distance from focal trees. Within 5-10 m away from a focal 302 tree, it was on average 29 times higher than expected by a completely random pattern, at 15-303 20 m it was still 12 times higher, but beyond 40 m it approximated the overall density λ (see 304 Appendix S1). As expected, the cumulative relative neighborhood density $K(r)/(\pi r^2)$ declined 305 somewhat slower and approximated the overall density λ at radius r of about 100 m (Fig. S3). 306 307 Thus, we expect the strongest density-dependent effects to occur within 20-30 m which should slowly fade away after 40-50 m (see Fedriani et al. 2010 for a more details on P. 308 bourgaeana clustering). 309

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311 *Fruit initiation*

Fruit initiation during 2012 and 2013 was low (12% and 8.5% on average) and highly variable 312 313 (range 0-40%) among P. bourgaeana trees during both years (Table 1). As expected, the rmark-correlation function for the percentage of flowers that initially set fruits during 2012 314 315 was significantly higher for individuals with conspecific neighbors within 35 m (Fig 1A; P =316 0.01). For example, a tree that had other trees at distance 25 m set about 1.4 times more fruit than the mean of all trees in the study area, and such differences declined quickly for above 317 35 m (Fig. 1A). Schlather's $I_{\rm mm}$ correlation function showed that fruit initiation of trees 318 separated by less than 20 m were positively correlated (Fig 1B; P = 0.015). The density 319 correlation function $C_{m,K}$ indicated a highly significant and positive density dependence at tree 320 neighborhoods > 5 m (P < 0.01; Fig. 1C) with a peak at 30m [$C_{m,K}(r) = 0.5$]. The 2013 data 321 exhibited very similar patterning, i.e. positive DDD of fruit initiation that was somewhat 322 stronger than that of the previous season (Fig. 1D-F). 323

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327 Fruit development

Fruit development rate from the total number of flowers that had set fruits (i.e. 440, 579, and 328 541 in 2011, 2012, and 2013, respectively) was over 3-fold higher than fruit initiation rate, but 329 330 also very variable among reproductive trees during the three seasons (Table 1). As predicted, and in contrast with the observed trend for fruit initiation, fruit development showed negative 331 density dependence during two of the three years. During 2011, fruit development was 332 significantly lower for individuals with neighbors within 50 m (Fig 2A; P < 0.01). For 333 example, the mean r-mark-correlation value for fruit development of trees with neighbors at 334 30 m was only about 0.54 times that of the population mean fruit development (Fig. 2A). The 335 spatial correlation in fruit development of two trees r apart was higher than expected for 336 individuals with conspecifics up to 50 m away (Fig 2B; P < 0.01). This was caused by the 337 consistently lower fruit development of nearby trees as shown by the r-mark correlation 338 function. As a consequence we observed strong negative density dependence with maximal 339 correlations between 40 and 80 m (P < 0.01) where trees within clusters showed lower than 340 341 expected fruit development (Fig. 2C).

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During 2012 the trend in fruit development was the same, being lower for *P. bourgaeana* individuals with close conspecific neighbors, though the observed values were not significant and fall within the simulation envelops (Fig 2D-F). During 2013, the r-mark correlation function indicated that fruit development was again significantly lower for individuals with conspecific neighbors up to scale of 40 m (P = 0.04; Fig. 2G). Also, fruit development exhibited a significant negative density dependence at neighborhoods with radius larger than 10m (P = 0.015; Fig. 2I).

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351 Overall fruit set

Overall fruit set of *P. bourgaeana* trees during 2012 and 2013 (0.054±0.006 and 0.028±0.007 352 353 fully developed fruits per flower, respectively) was generally low and variable among reproductive trees (Table 1). Interestingly, all three summary statistics during both seasons 354 355 were always non-significant and fall within the simulated envelops (always, P > 0.250; Fig. 3A-C [2012], Fig. 3 D-F [2013]). This result indicates lack of distance and density 356 dependence of overall fruit set as well as lack of spatial correlation among tree in overall fruit 357 358 set, thus the positive density dependence of fruit initiation was just neutralized by the negative 359 density dependence of fruit development.

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361 *Crop size and total seed number per tree*

Fruit crop size varied extensively among *P. bourgaeana* trees during the three seasons, 362 ranging over three orders of magnitude (Table 1). On average, trees produced twice as many 363 364 fruits during 2011 as during 2013 (Table 1). Crop sizes exhibited similar spatial patterning during 2011 and 2012, but differed in 2013 (Fig. 4) when fruit crops were smallest. During 365 2011 and 2012 individuals with nearby conspecific neighbors produced more fruits than 366 average (i.e., up to 30-40 m; Fig. 4A [2011], Fig. 4D [2012]), with differences being 367 significant in 2011 (P = 0.015). During 2011, fruit production was subject to significant 368 369 positive density dependence at scales greater than 10m (P < 0.01; Fig. 4C). During 2012, the pattern of density dependence was very similar to that in 2011, but somewhat weaker (P =370 0.015; Fig. 4F) and restricted to scales from 20-60 m. In contrast, during 2013, all three 371 summary statistics indicated no spatial structures in crop sizes at any smaller spatial scale (P 372 > 0.44; Fig. 4G-I). 373

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During 2011, the estimated total number of seeds per reproductive *P. bourgaeana* tree varied widely (i.e. up to two orders of magnitude; Table 1). However, all three summary statistics indicated a general lack of DDD effects on total number of seeds (P > 0.145; Fig. S4A-C).

These results indicate that the strong effect of conspecific neighbors at small scales for crop size (Fig. 4) did not influence the number of seeds per fruit. This is probably caused by trees with larger crops that produce fruits with fewer seeds, though such a trend was only marginally significant ($r_s = -0.207$, P = 0.096).

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383 DISCUSSION

384 We used detailed data from three consecutive seasons to examine DDD effects on sequential PRS components over a range of spatial scales. Application of the framework of marked point 385 pattern analysis allowed us to answer our three guiding questions. First, fruit initiation, fruit 386 387 development and crop size showed DDD effects over spatial scales that correspond to that of the highest neighborhood density of P. bourgaeana trees (some 30-40m, somewhat smaller 388 389 than the diameter of the typical cluster which yields 44m; Fedriani et al. 2010) as indicated by 390 the maximal correlation between reproductive success and neighborhood density. Second, we found the surprising result that the net effect of DDD on the overall fruit set of P. bourgaeana 391 392 trees just canceled because DDD effects on its two components (i.e., fruit initiation by fruit development) acted on opposite directions and with similar strengths. Finally, though the 393 DDD effects of different reproductive components of P. bourgaeana were conflicting in sign 394 395 (fruit initiation and crop size showed positive DDD effects whereas fruit development showed negative DDD effects), they showed similar critical scales imprinted by the strong clustering 396 of the trees. However, DDD effects did not appear every year. 397

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399 *Conflicting DDD processes*

400 Density and distance dependence in demographic parameters may be critical for the

401 regulation of many plant populations and thus for ecosystem functioning (Courchamp et al.

402 2008 review). Few studies, however, have considered DDD effects on several fitness

403 components simultaneously over a range of spatial scales which, as we illustrate, is needed

given that both sort of effects can emerge at contrasting plant life stages and spatial scales 404 405 (Aizen & Harder 2007). Our results provided clear evidence that different DDD processes affect different aspects of fruit production (Aizen & Vazquez 2007, Gómez et al. 2009, Jones 406 407 & Comita 2010). First, as expected, P. bourgaeana individuals with nearby neighbors (up to 40 m) initiated more fruits during both seasons studied, which is consistent with results from 408 409 other temperate and tropical species, including some hermaphrodite species (e.g., Kunin 1993, Aizen 1997; but see Spigler & Chang 2008). Such findings are generally attributed to the 410 attraction of pollinators to larger floral displays, and increased pollen transfer per visit at high 411 conspecific densities (Ishihama et al. 2006, Aizen & Harder 2007). 412

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However, fruit development showed negative density and distance dependence during two out 414 of the three monitored seasons (Fig. 2). One hypothesis to explain this result is that most 415 416 mating occurred between highly related P. bourgaeana individuals (e.g., Ishihama et al. 2006, Jones & Comita 2008). In entomophilous species pollen deposition from genetically related 417 418 donor conspecifics can strongly contribute to qualitative pollen limitation given that insect 419 pollinators typically disperse pollen locally (Harder & Barrett 1996, Aizen & Harder 2007). Though the initiation of fruit set can occur despite low-quality pollen, such developing fruits 420 421 are likely to abort because embryos homozygous for deleterious alleles tend to die during development (Herlihy & Eckert 2004, Aizen & Harder 2007). Several lines of evidence 422 support this hypothesis in our system. Genetic analyses indicates strong relatedness among 423 individuals within clusters (Rodríguez 2009), which is consistent with limited seed dispersal 424 425 and spatial contagion of dispersed P. bourgaeana seeds (Fedriani et al. 2010, Fedriani & Wiegand 2014). Such local genetic structure also occurs in other tree species experiencing 426 427 limited seed dispersal (Dutech et al. 2002, Ng et al. 2004, Voigt et al. 2009). Moreover, thorough hand-pollination experiments revealed that P. bourgaeana flowers supplemented 428 with pollen from individuals away from the focal cluster produced more fruits than 429

individuals supplemented with pollen from neighboring conspecifics (Fedriani et al. 2012).
Comparing levels of inbreeding in fully developed versus aborted seeds is a pending relevant
task (Jones & Comita 2008).

433

434 Further influential factors

Other factors, such as availability of soil resources, intraspecific competition, and tree size 435 436 could potentially affect fruit initiation and development, and hence overall fruit production (Pacala & Silander 1985, Obeso 2002, García-Camacho et al. 2009). However, we did not 437 find any effect of tree size (i.e. dbh and crown area) on fruit production and the locally 438 439 dominant sandy substrate appears homogeneous across our study area (Fedriani & Delibes 2009). Furthermore, by sampling all reproductive individuals within our study plot and 440 performing detailed analyses at a range of spatial scales, we have minimized any potential 441 442 effect of spatial heterogeneity, thus providing robust results. On the other hand, though it cannot be ruled out that high competition for resources and pollinators in high density 443 444 neighborhoods lessened tree reproductive success, high pollinator visitation and relatively high initial fruit sets and crop sizes are often observed in these tree clusters. Nonetheless, at 445 least during 2011, although individuals in large clusters produced more fruit (Fig. 4) their 446 447 seed production (a more pertinent aspect of female reproductive success) did not show such marked positive DDD pattern. This result indicates that although fruit production may 448 depend on flower production, subsequent conflicting DDD processes (including intraspecific 449 competition) can modify the spatial pattern of tree reproductive investment. 450

451

452 Conclusions

Our study revealed that complex scale-dependent processes operate at different components
of pollination and fruiting success of the self-incompatible Iberian pear *Pyrus bourgaeana* in
a Mediterranean area of southern Spain. Interestingly, DDD processes for sequential tree

reproductive success components operated in opposition and neutralized each other. This
result may help to explain why some authors failed to find expected DDD of plant
reproductive performance (see Ghazoul 2005 review). Understanding of the complex spatial
effects of DDD processes requires thus dissection of component processes to obtain the
complete picture since contrasting DDD processes may be hidden behind a single cumulative
measure of reproductive success.

462

Our approach can be widely applied to mapped plant populations for which detailed data on 463 reproductive success are available or attainable. Furthermore, point pattern analyses similar 464 465 to those presented here could be conducted using reproductive success components concerning subsequent plant stages (e.g. dispersed seeds, seedlings, saplings; Fedriani & 466 Wiegand 2014, Jacquemyn et al. 2010, Wiegand & Moloney 2014) and thus foster 467 understanding of population regulation. Because DDD effects concern many active areas of 468 ecological and evolutionary research (e.g., humanization, invasive species, density-dependent 469 470 selection; Courchamp et al. 2008, Thompson 2013), the combination of novel and classic mark correlation functions presented here could be broadly applied to unravel conflicting 471 mechanisms regulating the persistence of sessile organisms at a range of spatial scales. 472 473

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484 We do not perceive any potential source of conflict of interest influencing our objectivity in

the current investigation.

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487 REFERENCES

Aizen, M.A. & Vázquez, D.P. (2006) Flower performance in human-altered habitats. In: L.D
Harder y S.C.H. Barrett (editors). Ecology and Evolution of flowers. Pp. 159-179.

490 Oxford University Press.

- 491 Aizen MA & Harder (2007). Expanding the limits of the pollen-limitation concept: effects of
 492 pollen quantity and quality. Ecology, 88, 271-81.
- Aldasoro, J. J., Aedo C., & Muñoz-Garmendia F. (1996). The genus *Pyrus* L. (Rosaceae) in
 south-west Europe and North Africa. Biological Journal of Linnean Society, 121,
 143-158.

496 Augspurger C. K. (1981). Reproductive synchrony of a tropical shrub: experimental studies

- 497 on effects of pollinators and seed predators on *Hybanthus prunifolius* (Violaceae).
 498 Ecology, 62, 775-788.
- Baddeley, A., Diggle, P. J., Hardegen, A., Lawrence, T., Milne, R. K., & Nair, G. (2014). On
 tests of spatial pattern based on simulation envelopes. Ecological Monographs, 84,
 477-489.
- 502 Condit, R. et al. (2000). Spatial patterns in the distribution of tropical tree species. Science,
 503 288, 1414-1418.
- 504 Courchamp, F., Berec, L. & Gascoigne, J. (2008) Allee Effects in Ecology and Conservation.
 505 OxfordUniversity Press, Oxford.

- Dutech C. J. Seiter P. Petronelli H. I. Joly P. Jarne (2002). Evidence of low gene flow in a
 neotropical tree species in two forest stands of French Guiana. Molecular Ecology, 11,
 725-738.
- Fedriani, J. M., & M. Delibes (2009). Seed dispersal in the Iberian pear *Pyrus bourgaeana*: a
 role for infrequent mutualists. Ecoscience, 16, 211-221.
- Fedriani, J.M., Wiegand T, & Delibes, M. (2010). Spatial patterns of adult trees and the
 mammal-generated seed rain in the Iberian pear. Ecography, 33, 545-555.
- 513 Fedriani JM, G Calvo, M Zywiec, M Delibes (2012). Low pollination success in a small
- 514 population of the Iberian pear *Pyrus bourgaeana*: Allee effect or endogamic
- 515 depression? Ecological Society of Germany, Austria and Switzerland. 42nd Annual
- 516 Meeting GfÖ. P. 260.
- Fedriani J.M. & T. Wiegand. (2014). Hierarchical mechanisms of spatially contagious seed
 dispersal in complex seed-disperser networks. Ecology, 95, 514–526.
- 519 Fox, L.R. (2007). Climatic and biotic stochasticity: disparate causes of convergent
- demographies in rare, sympatric plants. Conservation Biology, 21: 1556–1561.
- 521 García-Camacho R, Méndez M, & Escudero A (2009) Pollination context effects in the high-
- 522 mountain dimorphic Armeria caespitosa (Plumbaginaceae): neighborhood is
- something more than density. Am J Bot, 96, 1620–1626.
- Getzin, S., M. Worbes, T. Wiegand & K. Wiegand. (2011). Size dominance regulates tree
 spacing more than competition within height classes in tropical Cameroon. Journal of
 Tropical Ecology, 27, 93-102.
- 527 Ghazoul, J. (2005). Pollen and seed dispersal among dispersed plants. Biological Reviews, 80,
 528 413-443.
- Gómez JM, Abdelaziz M, Camacho JPM, Muñoz-Pajares J, Perfectti F. (2009). Local
 adaptation and maladaptation to pollinators in a generalist geographic mosaic. Ecology
- 531 Letters, 12, 672–682.

533	experimental and natural plant populations. Journal of Ecology, 97, 567-580.			
534	Harder LD, Barrett SCH (1996) Pollen dispersal and mating patterns in animal-pollinated			
535	plants. In: Lloyd DG, Barrett SCH (eds) Floral biology: studies on floral evolution in			
536	animalpollinated plants. Chapman & Hall, New York, pp 140–190.			
537	Harper, J.L. (1977) Population biology of plants. Academic Press: New York.			
538	Herlihy CR & Eckert CG (2004). Experimental dissection of inbreeding and its adaptive			
539	significance in a flowering plants Aqilegia canadensis (Ranuculaceae). Evolution, 58,			
540	2693–2703.			
541	Herrera C (2009) Multiplicity in unity. Plant subindividual variation and interaction with			
542	animals. University of Chicago Press, Chicago.			
543	Hixon, M.A., S. W. Pacala, & S.A. Sandin. (2002). Population regulation: historical context			
544	and contemporary challenges of open vs. closed systems. Ecology, 83,1490-1508.			
545	Illian, J. et al. (2008). Statistical analysis and modelling of spatial point patterns. Wiley.			
546	Ishihama, F., S. Ueno, Y. Tsumura, & I. Washitani. (2006) Effects of density and fl oral			
547	morph on pollen fl ow and seed production of an endangered heterostylous herb,			
548	Primula sieboldii. Journal of Ecology, 94, 846 – 855.			
549	Jacquemyn, H., P. Endels, O. Honnay, & T. Wiegand. (2010). Spatio-temporal analysis of			
550	seedling recruitment, mortality and persistence into later life stages in the rare Primula			
551	vulgaris. Journal of Applied Ecology, 47,431–440.			
552	Janzen, D. H. (1970) Herbivores and the number of tree species in tropical forests. The			
553	American Naturalist, 104, 501-528.			
554	Jones FA, & Comita LS (2008) Neighbourhood density and genetic relatedness interact to			
555	determine fruit set and abortion rates in a continuous tropical tree population. Proc R			
556	Soc Lond B, 275, 2759–2767			
	23			

Gunton , R. M. , & W. E. Kunin. (2009). Density-dependence at multiple scales in

- Jones FA, & Comita LS (2010) Density-dependent pre-dispersal seed predation and fruit set
 in a tropical tree. Oikos, 119,1841–1847.
- 559 Knight, T. M., J. A. Steets, J. C. Vamosi, S. J. Mazer, M. Burd, D. R. Campbell, M. R.
- Dudash, et al. (2005). Pollen limitation of plant reproduction: Pattern and process.
 Annual Review of Ecology Evolution and Systematics, 36, 467 497.
- Kolb, A., R. Leimu & J. Ehrlén. (2007). Environmental context influences the outcome of a
 plant-seed predator interaction. Oikos, 116, 864-872.
- Levin, S.A. (1992). The problem of pattern and scale in ecology. Ecology, 73, 1943-67.
- 565 Mezquida ET & JM Olano (2013). What makes a good neighborhood? Interaction of spatial
- scale and fruit density in the predator satiation dynamics of a masting juniper tree.
- 567 Oecologia, 173, 483-492.
- Ng KKS, Lee SL, & Koh CL (2004) Spatial structure and genetic diversity of two tropical
 tree species with contrasting breeding systems and different ploidy levels. Molecular
 Ecology, 13, 657-669.
- 571 Nielsen, A. & Ims R. A., (2000). Bumble bee pollination of the Sticky Catchfly in a
- 572 fragmented agricultural landscape. Ecoscience, 7, 157-165.
- 573 Obeso, J. R. (2002). The costs of reproduction in plants. New Phytologist, 155, 321-348.
- 574 Pacala, S.W. & J.A. Silander, Jr. (1985). Neighborhood models of plant population dynamics.
 575 I. Single-species models of annuals. Amer. Natur., 125, 385-411.
- 576 Peterson, G. et al. (1998). Ecological resilience, biodiversity, and scale. Ecosystems, 1, 1432577 1435.
- 578 Raventós, J., T. Wiegand, & M. De Luis. (2010). Evidence for the spatial segregation
- 579 hypothesis: a test with nine-year survivorship data in a Mediterranean fire-prone
- shrubland show that interspecific and density-dependent spatial interactions dominate.
- 581 Ecology. 91:2110-2120.

- Rodríguez A. (2009). Clonal diversity and genetic diferentiation in *Pyrus bourgaeana* Decne.
 Ms. Thesis. University Pablo Olavide, Seville, Spain.
- Roughgarden, J., Gaines, S., & Possingham, H. (1988). Recruitment dynamics in complex life
 cycles. Science, 241: 1460-1466.
- Schlather, M., Ribeiro, P. J., Jr. & Diggle, P. J. (2004), Detecting Dependence between Marks
 and Locations of Marked Point Processes, Journal of the Royal Statistical Society, ser.
 B, 66, 79-93.
- Shimatani, K., (2002). Point processes for fine-scale spatial genetics and molecular ecology.
 Biom. J., 44, 325-352.
- 591 Strong, D.R. (1986). Density-vague population change. Trends Ecol and Evol. 1: 39-42.
- 592 Thompson, J. N. (2013). Relentless evolution. University of Chicago Press. Voigt, F. A. et al.
- 593 (2009). Linking seed dispersal and genetic structure of trees: a biogeographical
 594 approach. J. Biogeogr., 36, 242–254.
- Wiegand, T. et al. (2007). Analyzing the spatial structure of a Sri Lankan tree species with
 multiple scales of clustering. Ecology, 88, 3088 3102.
- 597 Wiegand, T, F. He, & S.P. Hubbell. (2013). A systematic comparison of summary
- characteristics for quantifying point patterns in ecology. Ecography. 36, 92-103.
- 599 Wiegand, T. & Moloney, K.A. (2014). Handbook of spatial point-pattern analysis in ecology.
- 600 Chapman & Hall / CRDC applied environmental statistics CRC Press / Taylor &
 601 Francis, Boca Raton, FL, 538 pp.
- 602 Żywiec M, M. Delibes, & JM Fedriani. (2012). Microgeographical, inter-individual, and
- 603 intra-individual variation in the flower characters of Iberian pear *Pyrus bourgaeana*
- 604 (Rosaceae) Oecologia, 169, 713-722.
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607	Supporting	Information
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609	Additional supporting information may be found in the online version of this article:			
610				
611	Appendix S1. Test of the density correlation function with artificial data			
612	Figure S1. Analysis of the spatial pattern of adult <i>P. bourgaeana</i> trees.			
613	Figure S2. The mark correlation functions for total seed number produced by target <i>Pyrus</i>			
614	bourgaeana trees during 2011			
615				
616	As a service to our authors and readers, this journal provides supporting information supplied			

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Table 1: Sample sizes, mean, standard error, minimum, and maximum values of different *Pyrus bourgaeana* reproductive success components used in our marked point pattern spatial analyses for each of the three years (2011, 2012, and 2013) encompassed during the study. N₁, summation of flower buds, seeds, or fruits monitored to estimate fruit initiation and fruit development, as well as total seed numbers. Note that fruit set is given as a proportion. N₂, number of reproductive trees considered for the estimation of each reproductive success component.

	N ₁	N ₂	Mean	SE	Minimum	Maximum
2011						
Fruit initiation	-					
Fruit development	440	71	0.595	0.047	0	1
Overall fruit set	-					
Crop size	36215	71	510.07	57.69	16	1858
Total seed number ⁺	78064	63	1239.1	188.2	59	7263
2012						
Fruit initiation	5012	67	0.120	0.011	0	0.345
Fruit development	579	61	0.455	0.046	0	1
Overall fruit set	5012	67	0.054	0.006	0	0.190
Crop size	17434	67	257.75	37.76	0	1448
2013						
Fruit initiation	5048	67	0.085	0.011	0	0.399
Fruit development	541	50	0.276	0.044	0	1
Overall fruit set	5048	67	0.028	0.007	0	0.351
Crop size	15725	72	218.40	30.58	0	1249

627

628 ⁺Estimated only during the 2011 fruiting season.

630 FIGURE LEGENDS

Figure 1: Fruit initiation. The three mark correlation functions (r-mark correlation function, 631 Schlather's I function, and density correlation function) to detect a potential spatial structure 632 in P. bourgaeana fruit initiation during 2012 (A-C) and 2013 (D-F). The r-mark correlation 633 function is the mean reproductive success component (i.e., fruit initiation) m_i of a tree at 634 635 distance r of another tree, Schlather's I_{mm} correlation function is the correlation between the fruit initiation m_i and m_i of two trees i and j separated by distance r, respectively, and the 636 neighborhood correlation function $C_{mK}(r)$ is the correlation between the mark m_i and the 637 number of neighbors within distance r. The observed mark correlation functions (closed 638 disks), the expected function under the null model of random mark i (grey solid line), and the 639 640 corresponding simulation envelopes (solid lines), being the 5th lowest and highest values of the functions created by 199 simulations under the null model, are shown. To test the overall 641 fit of the random marking null model over a range of distances we used Goodness-of-Fit 642 643 (Baddeley et al. 2014) for the entire distance interval up to 50m. The GoF test returns a Pvalue (shown in each panel) that indicates significant departures of the observed mark 644 correlation function from the random marking null model over the distance interval of 645 interest. 646

647

Figure 2: The mark correlation functions for fruit development in *Pyrus bourgaeana* during
2011 (A-C), 2012 (D-F), and 2013 (G-I). Other conventions are as described in Figure 1.

650

Figure 3: The mark correlation functions for the overall fruit set (i.e. the product of fruit
initiation by fruit development) in *Pyrus bourgaeana* during 2012 (A-C) and 2012 (D-F).
Other conventions are as described in Figure 1.

- **Figure 4**: The mark correlation functions for crop size in *Pyrus bourgaeana* during 2011 (A-
- C), 2012 (D-F), and 2013 (G-I). Other conventions are as described in Figure 1.



Figure 1.



Figure 2.



Figure 3



Figure 4