

1	Effects of climate on reproductive investment in a masting species: assessment of
2	climatic predictors and underlying mechanisms
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4	Running title: Climatic predictors of masting
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32 Summary

33 **1.** Mechanisms by which climatic factors drive reproductive investment and phenology in masting species are not completely understood. Climatic conditions may act as a 34 proximate cue, stimulating the onset of reproduction and indirectly increasing fitness 35 through benefits associated with synchronous reproduction among individuals. 36 Alternatively, climatic conditions may directly influence individual level allocation to 37 reproduction and reproductive success through effects occurring independently of 38 synchronous reproduction. We previously showed that masting in a ponderosa pine 39 (Pinus ponderosa) population was strongly influenced by spring mean temperature two 40 41 years before seed cone maturation (T_{i-2}) . However, recent work shows that the difference in temperature between previous growing seasons (ΔT) is more predictive of 42 reproductive investment in long-lived tree species. 43 44 2. Here we compared four candidate models that predict seed cone production in *P*. *ponderosa* based upon different climatic factors (including T_{i-2} and ΔT models). After 45 determining the best climatic predictor, we tested for a potential mechanism by which 46 climate might directly influence seed cone production independent of benefits via 47 synchrony, namely effects of temperature on trade-offs between current and past 48 49 reproduction (determined by underlying resource availability). 50 **3.** We found that T_{i-2} (rather than ΔT) was the best predictor of seed cone production. We further show that this same climatic factor exerts a direct fitness benefit to 51 52 individuals by reducing the strength of trade-offs between current and past reproductive efforts. 53 4. Synthesis: We demonstrate that a single climatic factor provides fitness benefits to 54 individuals directly, by weakening reproductive trade-offs, and indirectly through the 55 benefits associated with synchrony and masting. This suggests a mechanism for the 56

origin and maintenance of masting: individuals initially respond to climatic cues that
directly enhance reproduction (e.g. lower reproductive costs through weakened tradeoffs) and this dynamic, expressed across multiple individuals, reinforces these benefits
through the economies of scale associated with synchrony and masting.

61

Key-words: Economies of scale, *Pinus ponderosa*, Plant population and community
 dynamics, reproductive trade-offs, resource limitation, seed cone production, ΔT model

65 **INTRODUCTION**

66 Masting is a reproductive strategy defined as the episodic production of large,

67 synchronous seed crops by a plant population (Silvertown 1980; Kelly 1994). Several

68 fitness advantages have been associated with masting that entail economies of scale

such as seed predator satiation (Janzen 1971; Kelly et al. 2000; Fletcher et al. 2010),

70 improved seed dispersal (Norton & Kelly 1988; Kelly 1994), and increased pollination

efficiency (Kelly, Hart & Allen 2001; Kon *et al.* 2005a; Rapp, McIntire & Crone 2013;

72 Moreira *et al.* 2014), all of which have been invoked to explain the occurrence and

73 maintenance of this reproductive phenomenon.

The reproductive dynamics of masting species are strongly influenced by climatic conditions previous to or during the initiation of reproduction, which influence allocation to reproduction and the degree of synchrony in seed production (Schauber *et al.* 2002; Kon *et al.* 2005b; Kelly *et al.* 2008; Smaill *et al.* 2011; Roland, Schmidt & Johnstone 2014). For example, high seed production is frequently linked to warm temperatures during the previous growing season (e.g. Schauber *et al.* 2002; Selås *et al.* 2002; Kelly *et al.* 2008; Masaki *et al.* 2008). In addition, other studies have found that

drought in the early summer has positive effects on flowering the following year (e.g.
Piovesan & Adams 2001; Krebs *et al.* 2012).

The mechanisms by which climatic factors drive reproductive patterns in 83 masting species are not fully understood, but several non-mutually-exclusive scenarios 84 are commonly considered. First, climatic conditions may act as proximate cues that 85 directly stimulate the onset of reproduction and reproductive phenology (Kelly & Sork 86 87 2002; Kelly et al. 2013). Under this scenario, the fitness advantage of individual plants responding similarly (in amount and timing) originates from the economies of scale 88 associated with population-level reproductive synchrony (see above; Norton & Kelly 89 90 1988; Kelly 1994). Nevertheless, it is difficult to explain the evolution of a reproductive 91 response to a cue that provides no direct, individual-level benefit (Koenig et al. 2015). 92 Second, other authors argue that climatic factors shape reproductive patterns at the 93 individual level, and that any benefit of masting associated with economies of scale arises secondarily. Climatic factors may directly alter resource availability (Koenig et 94 95 al. 1996; Richardson et al. 2005; Pearse, Koenig & Knops 2014; Koenig et al. 2015) such that there are individual-level fitness advantages of responding to climatic factors 96 97 (Pearse, Koenig & Knops 2014; Koenig et al. 2015). For example, seasonality in 98 precipitation or temperature have been shown to influence plant nitrogen and carbon availability (Allen & Platt 1990; Richardson et al. 2005; Smaill et al. 2011), as well as 99 soil moisture (Abrahamson & Layne 2003; Richardson et al. 2005) which may in turn 100 101 influence within-plant resource allocation patterns (Barringer, Koenig & Knops 2013; Pearse, Koenig & Knops 2014). As a consequence of such effects on resource 102 103 availability, climatic conditions may strengthen or weaken trade-offs between current and past reproduction; such effects may be particularly strong in the case of masting 104 species due to resource depletion during large reproductive events (Sala et al. 2012). 105

Finally, so called "pollen coupling" has been proposed to underlie masting in the
absence of climatic drivers (see reviews by Miyazaki 2013; Crone & Rapp 2014), where
inter-annual variation is driven by reproduction trade-offs among years, and synchrony
occurs because out-cross pollen limitation renders asynchronous reproduction
unsuccessful.

A recent analysis of long-term datasets of 15 species across five plant families 111 112 proposed a new predictive model to explain how climatic conditions modulate reproductive investment and phenology in masting species (Kelly et al. 2013). 113 Specifically, for 12 of the 15 species studied, a model using the difference in 114 115 temperature from one previous growing season to the next (ΔT model hereafter) as a 116 predictor of reproductive investment better predicted (i.e. smaller Akaike information 117 criterion and higher correlation coefficient) seed production relative to a model 118 including the temperature from a single previous growing season as predictor (Kelly et al. 2013). The authors offered several reasons for why the ΔT model exhibited a better 119 120 fit relative to the model based upon a temperature from a single previous season. First, a model based on temperature from a single previous season is not capable of explaining 121 122 why two previous consecutive warm years rarely cause two consecutive years of high 123 seed production (i.e. climatic effects of previous consecutive years are not additive). 124 Second, the ΔT model predicts that long-term plant reproductive responses are not 125 influenced by gradual increases in mean temperature (such as those produced by climate 126 change) but rather by averaging effects of climatic conditions across multiple growing seasons. This would also explain why the frequency of high-seed years (i.e. mast years) 127 128 remains relatively constant over large periods of time (Kelly et al. 2013). However, not all studies have been supportive of this model's predictions. For example, Koenig & 129 Knops (2014) showed that acorn production of three out of four oak species was 130

131 correlated with spring and summer temperature and precipitation, but not with 132 differences between previous growing seasons for each of these climatic variables. 133 Accordingly, further tests are needed to determine the general applicability of the ΔT 134 model, as well as to identify the climatic drivers that best predict reproduction in long-135 lived species and the mechanisms underlying such effects.

Using a long-term data set (31 years), we previously reported that masting 136 137 behaviour in a population of ponderosa pine (*Pinus ponderosa* Douglas ex C. Lawson) increases reproduction through two complementary mechanisms operating via 138 synchrony and associated economies of scale. First, synchronized, high levels of 139 140 reproduction reduced pollen limitation by increasing the rate of female cone fertilization 141 (Moreira et al. 2014). Second, synchronous bouts of high cone production, followed by 142 intervening years of low reproduction reduced seed cone herbivory through predator 143 satiation (Linhart et al. 2014). Additionally, we found that production of mature seed cones in this population was strongly influenced by spring (May-July) mean 144 145 temperature two years before cone maturation (Ti-2 model hereafter) (Mooney, Linhart & Snyder 2011). 146

147 Here we sought to determine which climatic factors drive synchronous 148 reproduction in *P. ponderosa* by comparing different competing models, and address if responses to such climatic factors provide direct, individual-level reproductive benefits 149 150 independent of those previously shown to occur through synchrony (Mooney, Linhart & 151 Snyder 2011; Linhart et al. 2014; Moreira et al. 2014). First, we compared four candidate models that predict mature seed cone production in *P. ponderosa* based upon 152 153 different climatic factors: 1) a model using as predictor the difference between spring mean temperature two and three years before cone maturation (ΔT model), i.e. T_{i-2} - T_{i-3} 154 (Kelly et al. 2013), 2) a model using as predictor the mean temperature two years before 155

cone maturation (i.e. T_{i-2}; Mooney, Linhart & Snyder 2011), 3) a model using as 156 157 predictor the temperature three years before seed cone maturation (T_{i-3}) , and 4) a model using as predictor the individual effects of spring temperature two years and spring 158 temperature three years before cone maturation (2T model hereafter, i.e. T_{i-2} and T_{i-3}) 159 (Kelly et al. 2013). Second, after determining which temperature predictor best 160 explained seed cone production (based upon results from the above models), we tested 161 162 for a potential mechanism by which such climatic variables could provide a direct reproductive advantage not associated with the economies of scale from synchronized 163 reproduction. Specifically, we tested whether temperature influenced within-plant 164 165 resource allocation patterns via reproductive trade-offs. In so doing, this study identifies 166 the climatic factors influencing reproductive investment and phenology in a long-lived 167 plant and proposes a mechanism by which masting originates and is maintained.

168

169 MATERIAL AND METHODS

170 Pine species, study area, and data collection

171 Ponderosa pine is native to western North America, but has a widespread distribution as

a planted species throughout temperate areas of both the New and Old World

173 (Richardson 1998). Episodic mast seeding events are common for this species (e.g.

174 Mooney, Linhart & Snyder 2011; Linhart et al. 2014; Moreira et al. 2014). Female

175 cones ("seed cones" hereafter) are usually found on the upper branches and are

176 produced in early spring and require two growing seasons after pollination to mature,

177 reaching their full size by mid-summer. Several months later, during the fall, seed cones

178 open and release their seeds.

We carried out a long-term (31 years) field survey during which we monitoredthe reproduction of 217 individuals from a ponderosa pine population distributed over a

2-ha area on the south-facing slope of Boulder Canyon, at an elevation of 1740 m in the 181 182 Front Range of the Rocky Mountains (near the town of Boulder, Colorado, 40° 00' 48"N, 105° 18' 12"W; Linhart & Mitton 1985; Mooney, Linhart & Snyder 2011). To 183 reduce among-tree variation in abiotic factors strictly associated to fine-scale 184 185 environmental conditions at each tree location, we selected the experimental trees from an area with a uniform slope, soils, sun and wind exposure. At the beginning of this 186 187 study, the age of the studied trees ranged from ca. 40 to over 280 years based upon trunk cores taken at 30-50 cm above ground level. We also measured the basal diameter 188 of each tree at the beginning of this study. 189

190 We recorded seed cone production during each year (from July to October for 31

191 years [from 1977 to 2008, except in 2004]) by visually counting the total number of

192 mature seed cones (pollinated during the previous year and matured during the current

193 year) found throughout the crown of each tree (Mooney, Linhart & Snyder 2011;

194 Linhart et al. 2014; Moreira et al. 2014). Previous studies in conifers have documented

that the number of seed cones per tree is a good predictor (positive correlate) of the

number of seeds per tree (e.g. Zasada & Viereck 1970; El-Kassaby & Cook 1994; Krebs

et al. 2001). Finally, we also gathered climatic data (monthly mean temperature and

total monthly precipitation) for this population from the Colorado Climate Center at the

199 Department of Atmospheric Science of Colorado State University

200 (http://ccc.atmos.colostate.edu/, site = Boulder 50848) (Mooney, Linhart & Snyder

201 2011).

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203 Statistical analyses

204 Comparison of climatic predictors of ponderosa pine reproduction. In a previous study

using the same population, we related mean monthly precipitation and temperature to

mature seed cone production (Mooney, Linhart & Snyder 2011). We found that 206 207 increased mature seed cone production was associated with decreases in spring 208 temperature (mean of May, June, July) two years before, increases in summer precipitation (mean of July, August, September) two years before, and increases in 209 210 winter temperature (mean of December, January, February) during the previous year (Mooney, Linhart & Snyder 2011). To achieve normality of residuals in our models, 211 212 seed cone data were log-transformed for the present analyses (see below). After log-213 transforming, spring temperature two years before was the only significant predictor of 214 mature seed cone production. Additionally, because pollen and ovule meiosis 215 corresponded with spring temperature two years before seed cone maturation (Mooney, 216 Linhart & Snyder 2011), the effect of this climatic factor would likely be the best 217 predictor of reproductive investment. For both these reasons, in the present study we 218 only tested models that included predictors based upon spring temperatures (calculated as the mean of May, June and July monthly means; Mooney, Linhart & Synder 2011). 219 220 We compared four candidate models using log-transformed seed cone data at the population level (Kelly et al. 2013; Pearse, Koenig & Knops 2014). First, we assessed 221 222 the predictive power of the T_{i-2} model, where increased mature seed cone production 223 was previously shown to be associated with decreases in spring (May-July) temperature 224 two years before at the studied ponderosa pine population (Mooney, Linhart & Snyder 2011). Second, we sought to compare the T_{i-2} model with a model based on the 225 226 difference in temperature between T_{i-2} and the year preceding the initiation of reproduction (T_{i-2} - T_{i-3}), i.e. the ΔT model. Finally, for the sake of thoroughness, we ran 227 228 two additional models, one including both T_{i-2} and T_{i-3} (2T model) and the other including T_{i-3} alone (T_{i-3} model) (Pearse, Koenig & Knops 2014). The four candidate 229 230 models were compared using the Pearson r correlation coefficient and the corrected

Akaike information criterion (AIC_c) (see Krebs *et al.* 2012; Kelly *et al.* 2013; Pearse, Koenig & Knops 2014). The AIC_c is a measure of the relative fit of a statistical model based upon the observed data. The model with the smallest AIC_c has the best fit because it minimizes the information loss (Burnham & Anderson 2002; Murtaugh 2014). In addition, this information criterion penalizes the model based upon the number of parameters included (Burnham & Anderson 2002; Murtaugh 2014).

Mechanism of climatic effects on reproduction. Following Pearse, Koenig &
Knops (2014), we explored the mechanistic links between climatic factors and seed
cone production by testing whether temperature influenced patterns of reproductive
investment (via effects on resource availability and trade-offs). In so doing, we assessed
whether there were individual-level effects of climate on plant reproductive investment
that occurred independently of fitness benefits of population-level reproductive
synchrony.

In particular, we tested for a trade-off between current and past female 244 245 reproduction, and whether climatic factors influenced the strength of this trade-off. To 246 test for such trade-off, we regressed current mature seed cone production onto mature 247 seed cone production in the previous year. Then, to test for an effect of temperature on 248 this trade-off, we performed a linear mixed model with seed cone production in the 249 previous year, $T_{i,2}$ (the spring mean temperature two years before seed cone maturation), 250 and their interaction as predictors of mature seed cone production. Because reproductive 251 trade-offs occur at the individual level due to resource limitation and allocation constraints, this mechanistic model was conducted at the tree level (Pearse, Koenig & 252 253 Knops 2014). We used T_{i-2} in this model (instead of the other climatic predictors) because this climatic variable was the best predictor of mature seed cone production 254 (see *Results*). The interaction term tested whether the relationship between current and 255

past reproduction was contingent upon the spring mean temperature two years before 256 seed cone maturation (i.e. climate influencing reproductive investment via within-tree 257 258 resource availability). A negative value for the interaction parameter indicates that as T_i-2 increases, the relationship between current seed cone production and past seed cone 259 260 production becomes more negative (i.e. stronger trade-off between seed cone production in N_{i-1} and N_i), or alternatively, that as T_{i-2} decreases (cooler previous springs), the 261 262 trade-off becomes weaker. To account for effects of tree size which have previously been shown to influence the magnitude of reproductive trade-offs (i.e. larger trees have 263 more available resources and are less likely to exhibit allocation constraints), we 264 265 included basal tree diameter as a covariate in this statistical model (Almqvist, Jansson & 266 Sonesson 2001; Smaill et al. 2011; Santos-del-Blanco & Climent 2014). 267 The test of mechanism (trade-offs) of temperature effects on reproduction was

performed using linear mixed models with PROC MIXED in SAS 9.2 (SAS Institute,
Cary, NC), using tree as a random factor to account for repeated measures taken from
each tree throughout the sampling period (Moreira *et al.* 2014; Pearse, Koenig & Knops
2014).

272

273 **RESULTS**

A total of 194,052 seed cones were produced at the site over the 31 years of study for a site-wide average of $6,064 \pm 1,410$ cones per year (mean \pm SE). Seed cone production varied extensively among the 217 trees, ranging 0 to 26,040 seed cones produced throughout the 31 years sampled.

278

279 Models of climatic predictors of seed cone production in ponderosa pine

280 The ΔT , T_{i-2} and 2T models all significantly predicted mature seed cone production by 281 *P. ponderosa*, whereas the T_{i-3} model was not significant (Table 1). The Pearson r correlation coefficients between climatic factors and mature seed cone production were 282 negative and similar in magnitude for these three significant models (Table 1, Fig. 1), 283 284 demonstrating a pattern of increase in mature seed cone production with (i) decreases in spring mean temperature two years prior to cone production (T_{i-2} model), (*ii*) decreases 285 286 in the spring mean temperatures two and three years prior (T2 model), and (*iii*) 287 decreases in the change in spring mean temperatures from three years prior to two years prior (Δ T model). However, the T_{i-2} and 2T models had smaller AIC_c values and thus 288 289 provided a better fit relative to the ΔT model (Table 1, Fig 1). These results indicate that 290 the difference in temperature between previous growing seasons (as proposed by Kelly 291 et al. 2013) was a less robust predictor of mature seed cone production compared with 292 temperature two years before seed cone maturation (Ti-2) or temperature two years before together with three years before maturation (2T). In addition, although the T_{i-2} 293 294 and 2T models had qualitatively similar AIC_c values, we use T_{i-2} for subsequent analyses because this model had a marginally better fit (0.6 lower AIC_c). Moreover, 295 296 while the AIC accounts for the number of terms in a model, the T_{i-2} model provided a 297 more parsimonious explanation of mature seed cone production relative to the 2T model 298 as the former included only one predictor and the latter included two. $T_{i\text{-}2}$ ranged from 16.23 °C in 1995 to 21.62 °C in 2000. The mean $T_{i\text{-}2}$ over the 31 299

years of study was 18.76 ± 0.22 (mean \pm SE). Importantly, for climate to drive masting, patterns of reproduction must be more variable than climate (Kelly 1994). Accordingly, we found that the coefficient of variation of seed cone production at the population level was 1.26 (Linhart *et al.* 2014) while the coefficient of variation of T_{i-2} was 0.06.

304

Temperature as a driver of seed cone production in ponderosa pine

306 In accordance with the expectation of reproductive trade-offs, we found that the current 307 year's mature seed cone production (N_i) was negatively affected by the production of mature seed cones during the previous year (N_{i-1}) (Table 2a). Moreover, we found a 308 significant interaction between seed cone production in the previous year and T_{i-2} on 309 current seed cone production (Table 2b), showing that the observed reproductive trade-310 311 off was contingent upon the spring temperature two years before seed cone maturation. A negative value for the interaction parameter was observed (interaction between T_{i-2} 312 and N_{i-1} ; Table 2b), which indicates that as T_{i-2} increases, the relationship between 313 314 current seed cone production and past seed cone production becomes more negative (i.e. 315 stronger trade-off between seed cone production in N_{i-1} and N_i), or alternatively, that as 316 T_{i-2} decreases the relationship becomes less negative (Fig. 2). This demonstrates that 317 cool spring temperatures two years previous to mature seed cone production are not only positively associated with masting (and the benefits obtained due to the economies 318 319 of scale from synchrony; Moreira et al. 2014; Linhart et al. 2014), but are also 320 associated with an independent, direct positive effect on reproduction by decreasing the 321 strength of individual-level trade-offs between current and past reproduction (Fig. 2). 322

323 **DISCUSSION**

324 <u>Overview</u>

Our past research with the same *P. ponderosa* population demonstrated reproductive benefits of masting through both predator satiation and pollination efficiency (Linhart *et al.* 2014; Moreira *et al.* 2014). In the present work, we further show that the same climatic conditions that drive masting exert direct effects on individuals by influencing the strength of trade-offs between current and past reproduction. This suggests that

individual-level responses to climatic factors have the potential to influence inter-annual
patterns in plant reproduction in the absence of the collective benefits associated with
masting. Together, these findings suggest a potential mechanism that explains the origin
and maintenance of masting: individuals initially respond to climatic cues that directly
enhance reproduction (e.g. lower reproductive costs through weakened trade-offs) and
this dynamic, expressed across multiple individuals, reinforces these benefits through
the economies of scale associated with synchrony and masting.

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341

338 <u>Models of climatic predictors of seed cone production</u>

339 Our findings run counter to the propposition by Kelly *et al.* (2013) that the Δ T model is 340 superior for predicting seed production than other climatic variables. They concluded

342 insensitive to increases in global mean temperature. However, to date only three studies

that the ΔT model was more robust to the introduction of additional data, as well as

343 spanning 20 masting species across six families have tested the relative fit of this model,

and their findings provide mixed support (Kelly *et al.* 2013; Koenig & Knops 2014;

Pearse, Koenig & Knops 2014). For example, a study by Kelly *et al.* (2013) was

generally supportive as they found that the ΔT model had a better fit in predicting seed

production for 12 of 15 studied plant species (see Table 2 in Kelly *et al.* 2013).

However, Koenig & Knops (2014) found that temperature and precipitation during the

349 previous spring and summer were strongly linked to acorn production by four oak

350 species (*Quercus* spp.), whereas a ΔT model failed to predict acorn production for three

of these species. Similarly, Koenig *et al.* (2015) documented that temperatures during

the spring flowering period (but not temperature difference) in *Quercus lobata* altered

the patterns of synchrony and temporal variability in acorn production. Finally, Pearse,

Koenig & Knops (2014) found that the ΔT model explained only a slightly higher

proportion of acorn production by *Q. lobata* than a model that included April temperature alone (i.e. similar R^2 and P-values).

Two factors may explain why our findings for ponderosa pine (and several other 357 studies) have failed to support the ΔT model. Specifically, (i) differences in 358 environmental heterogeneity and in the steepness of the environmental gradients among 359 species' distribution ranges (e.g. alpine grasslands and lowland forests in Kelly et al. 360 361 2013 vs. montane forests in our case), and (ii) differences in species life forms and longevity (herbs and small trees in Kelly et al. 2013 vs. a long-lived, large tree). Both 362 factors may result in species-to-species variability in the effects of temperature on 363 364 reproductive investment and phenology. In this context, pines (and oaks) are long-lived 365 species that occupy a wide range of edaphic and climatic conditions, spanning from 366 temperate to tropical regions of the world and from sea level to high elevations 367 (Richardson 1998). Therefore, it is likely that long-lived tree species with broad distributions (as opposed to short-lived herbs with narrower distributions) will exhibit a 368 369 greater range of responses to temperature variation and will be better adapted to 370 assimilate temperature data over long periods of time (as opposed to biannual cycles as 371 proposed by the ΔT model), and adjust their reproduction accordingly.

372

373 Drivers of seed cone production in ponderosa pine

The responses of co-occurring individuals to a single climatic cue may provide fitness

advantages to individuals through the economies of scale associated with synchrony

- 376 (Silvertown 1980; Kelly 1994). Our past studies of this population demonstrate
- 377 synchrony in response to cues that indirectly and positively affect individual fitness
- through both predator satiation (Linhart *et al.* 2014) and increased pollination efficiency
- 379 (Moreira *et al.* 2014). Such fitness benefits from synchrony are suggestive of natural

selection for response to a common cue, but it is difficult to explain the evolution of a
reproductive response to a cue that provides no direct, individual level benefit (Koenig *et al.* 2015).

For *P. ponderosa*, cool spring temperatures two years before seed cone 383 maturation reduced the strength of the trade-off between past and current reproduction 384 (Fig. 2), thus providing a direct reproductive benefit for individuals responding to this 385 386 cue. Our finding of reproductive trade-offs is similar to that reported for other longlived tree species (Sork, Bramble & Sexton 1993; Koenig et al. 1994; Crone, Miller & 387 Sala 2009). The prevailing interpretation of such patterns has been that the production 388 389 of a large seed crop depletes substantial amounts of stored resources, resulting in 390 allocation constraints during subsequent reproductive events (e.g. Crone, Miller & Sala 2009; Sala et al. 2012; Han et al. 2014). Accordingly, Kelly (1994) proposed that 391 392 individual variation in seed output in masting species might largely depend on how heavily each plant invests resources during masting events, the so-called "depletion 393 394 coefficient". If this depletion coefficient is high, plant reserves would be depleted and would not subsequently reproduce again for some time. 395

396 An alternative mechanism proposed to explain masting behaviour independently 397 of climatic cues is given by the "pollen coupling hypothesis". This hypothesis describes how endogenous resource dynamics linked to pollen limitation can drive masting in the 398 absence of climate drivers (see reviews by Miyazaki 2013; Crone & Rapp 2014). This 399 400 hypothesis assumes density-dependent pollination, a high cost of producing seed as compared to pollen and ovules, and reproductive trade-offs among years (Isagi et al. 401 402 1997). During years of low reproduction, seed set of reproducing individuals is low and stored resources are not depleted, promoting future reproductive investment. In contrast, 403 during years of higher reproduction, seed set is high and stored resources are depleted, 404

thus reducing future reproductive investment. Accordingly, over time individuals are 405 eventually entrained into synchronous and variable reproduction (i.e. masting) in the 406 407 absence of any climatic cues. In our ponderosa pine population, we demonstrated two of the conditions necessary for pollen coupling to function: pollen limitation (Moreira et 408 409 al. 2014) and a negative correlation between past and current reproduction (current 410 study). Accordingly, while this study was not aimed at testing this hypothesis, we 411 acknowledge that other mechanisms not associated with climatic cues might also be at 412 work and explain masting events in this population.

We propose a model for the ecological and evolutionary origins of masting 413 414 based upon our findings for ponderosa pine (Fig. 3). First, there is selection to time 415 reproduction in response to a climatic factor providing a direct fitness benefit to the 416 individual (i.e. selection for initiating reproduction during cool springs reduces 417 reproductive trade-offs and thus minimizes the costs of reproduction). Such responses, when expressed across multiple individuals, result in synchrony. This synchrony may in 418 419 turn be enhanced through two complementary mechanisms, one evolutionary and one 420 ecological. First, selection for response to the climatic cue may be strengthened by the 421 indirect fitness benefits associated with synchrony and economies of scale (e.g. 422 pollination efficiency, predator satiation). And second, in the absence of any additional 423 selection, synchrony may be strengthened through the ecological dynamics of pollen 424 coupling (see above). Finally, these mechanisms might complement each other through 425 feedbacks, where pollen coupling leverages a small evolutionary response into population-level synchrony, which could in turn reinforce selection. It is important to 426 427 note, however, that the interpretation of our results within an evolutionary context should be made with caution as we measured the reproductive response to a climate cue 428 429 on annual scales and throughout a portion of this species lifetime. In this sense, it is

difficult to demonstrate that a particular pattern of reproduction in a long-lived species
leads to greater life-time fitness (i.e. cumulative seed cone production over the life-time
of an individual).

433

434 <u>Future directions</u>

Because climatic variables correlated with seed cone production can dramatically vary 435 436 from site to site (e.g. Mooney, Linhart & Snyder 2011; Koenig & Knops 2014), the observed effects of climate on trade-offs between past and current reproduction may 437 therefore be contingent upon site-specific conditions (e.g. plant species, environmental 438 439 heterogeneity). Based on this, a combination of long-term data sets (such as that in this 440 study) collected at multiple sites is needed for more tree species in order to determine the relative importance of different climatic drivers influencing plant reproductive 441 442 investment and phenology, as well as to assess the general mechanisms underlying such effects. In addition, our results call for further studies and analyses that separate direct, 443 444 individual-level effects of climate on reproductive investment, from indirect, population-level effects via synchrony. In doing so, we will be able to establish a link 445 446 between individual-level dynamics (e.g. via trade-offs as in this study) and population-447 level dynamics via synchrony. Furthermore, by addressing specific plant traits underling 448 these individual-level responses as well as selection upon such traits we will be able to understand the mechanisms determining the evolution of masting. Finally, we call for 449 450 comparative studies across habitat types, plant growth forms, and life histories in order to shed light into the relative influence of exogenous (abiotic factors) and endogenous 451 452 (resource allocation patterns) processes driving long-term patterns of reproductive investment and phenology. 453

454

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- 463

464 Data accessibility

- 465 Data are available in the Dryad Digital Repository, doi: 10.5061/dryad.h29r7 (Moreira
- *et al.* 2015). Data will be publicly available 12 months after publication.

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602	Table 1. Climatic predictors of mature seed cone production in <i>Pinus ponderosa</i> . Four
603	candidate models were evaluated: T_{i-2} model (spring mean temperature two years before
604	mature seed cone production), T_{i-3} model (spring mean temperature three years before
605	mature seed cone production), ΔT model (change in spring mean temperature from
606	three to two years before seed production, T_{i-2} - T_{i-3}), and 2T model (individual effects
607	of spring mean temperature three and two years before seed production). Corrected
608	Akaike information criterion (AIC _c), Pearson r correlation coefficients, and P -values are
609	shown. Significant predictors ($P < 0.05$) are typed in bold.

T _{i-2}	95.4		
_	2011	-0.605	<0.001
T _{i-3}	108.5	0.081	0.664
ΔΤ	108.9	-0.517	0.002
2T (T _{i-2} , T _{i-3})	96.0	-0.615	0.001

623	Table 2. Results from models linking mature seed cone production (N_i) and resource
624	limitation in <i>Pinus ponderosa</i> . (a) Regression model testing for a relationship between
625	current mature seed cone production and mature seed cone production during the
626	previous year (N_{i-1}) , i.e. test of reproductive trade-off. (b) Linear mixed model testing
627	for the effects of mature seed cone production during the previous year (N _{i-1}), spring
628	mean temperature two years before seed cone maturation (T_{i-2}), and the interaction
629	between N_{i-1} and T_{i-2} (i.e. effect of temperature on the trade-off). Statistical analyses of
630	both mechanistic models were performed at the individual tree level using a mixed
631	model with tree as a random factor and tree basal diameter at the beginning of this study
632	(D) as a covariate. The slope estimator (β) with the standard error (inside brackets), F-
633	values with the degrees of freedom (inside brackets), and <i>P</i> -values are shown.
634	Significant predictors ($P < 0.05$) are typed in bold.

	a)	$N_i = N_{i\text{-}1} + D$	l i i i i i i i i i i i i i i i i i i i	b) $N_i = N_i$	$I_{i-1} + D + T_{i-2} + D$	+ $N_{i-1} \times T_{i-2}$
	β	F-value	Р	β	F-value	Р
	(s.e.)	(df)		(s.e.)	(df)	
N _{i-1}	-0.0943	58.55	<0.001	0.6075	19.60	<0.001
	(0.0123)	(1, 5866)		(0.1372)	(1,5864)	
Diameter (D)	0.3119	130.58	<0.001	0.3084	130.68	<0.001
	(0.0273)	(1,5866)		(0.0270)	(1,5864)	
T _{i-2}	-	-	-	-14.699	322.87	<0.001
				(0.818)	(1,5864)	
$N_{i\text{-}1}\!\times T_{i\text{-}2}$	-	-	-	-0.0369	26.22	<0.001
				(0.0072)	(1,5864)	

FIGURE LEGENDS

Figure 1. Climatic predictors of seed cone production. Relationships between log-transformed seed cone production in *Pinus ponderosa* and (a) spring temperature (May-July) from two years before seed production (T_{i-2}) and (b) the difference in mean spring temperature two and three years before mature seed production (ΔT). Each point represents a year (N = 31). Pearson r correlation coefficients and P-values are shown. Figure 2. Model predictions for the relationship between current seed cone production (N_i) and past seed cone production (N_{i-1}) at spring temperatures (mean of May, June and July two years before seed production; T_{i-2}) representing the maximum, minimum, and averaged temperatures observed for the studied population of ponderosa pine. Figure 3. Diagram representing proposed links between individual- and population-level responses to climatic factors for the studied ponderosa pine population. Pollen coupling was included as it has been proposed to underlie masting in the absence of climatic drivers, where inter-annual variation is driven by reproductive trade-offs among years, and synchrony occurs because out-cross pollen limitation renders asynchronous reproduction unsuccessful. ⁽¹⁾ Mooney et al. 2011; ⁽²⁾ this study; ⁽³⁾ Moreira et al. 2014; ⁽⁴⁾ Linhart et al. 2014.











Previous seed cone production (N_{i-1})

- **Figure 2.** Moreira et al





