1   **Meta-analysis of meta-analyses in plant evolutionary ecology**

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23   **Abstract.** After two decades of meta-analyses on plant traits, we can now look for

24   global emergent patterns in plant evolutionary ecology. Hundreds of meta-analyses have

25   focused on the effects of specific selection pressures on plant fitness, and the buildup of

26   such results allows us to ask general questions regarding selection pressures and plant

27   responses, a major focus of evolutionary ecology. Plant traits are affected by both

28   abiotic and biotic factors. For example, biotic pressures like herbivory may affect

29   physiological (i.e. secondary defences) and reproductive (i.e. seed predation) traits.

30   Similarly, abiotic pressures such as increased CO_2 may affect both plant physiology and

31   reproduction. We tested whether biotic or abiotic selective pressures are more important

32   for plant traits, and if the strength of the response to those pressures depends on the

33   plant trait studied by meta-analyzing published meta-analyses on plant responses. We

34   classify meta-analyses according to the type of response variable studied (fitness and

35   non-fitness traits) and the type of selective pressure examined (biotic or abiotic). Our

36   database showed biases in the meta-analysis literature, for example that the majority of

37   studies are focused on non-fitness traits, i.e. on traits that are not directly related to

38   reproduction or survival, and furthermore, on non-fitness traits under abiotic selection

39   pressures. The meta-meta-analysis showed that the strength of responses to selection

40   depends on the nature of selection (stronger for biotic than for abiotic factors) but,

41   unexpectedly, not on the type of trait under study as previously found. The stronger

42   responses to biotic factors can be explained if biotic selection is more variable in space

43   and time, driven by interactions with other organisms. The relative importance of biotic

44   versus abiotic factors on plant traits has been little studied in the past, and would benefit

45   from more studies and reviews that fill the under-represented combinations of selective

46   pressures and plant traits (i.e. abiotic factors on fitness traits).

Introduction

After two decades of meta-analyses on plant traits, beginning with the introduction of the technique in the field of ecology in the early 1990's (see Gurevitch et al. 2001), it is timely to look for global emerging patterns in the literature. In the field of plant evolutionary ecology specifically, hundreds of meta-analytical studies have focused on the effects of specific selection pressures on a single or a few plant fitness components. As a result, generalizations can be made, for example on how increased N availability or increased herbivore activity can affect plant growth or reproductive output and how such effects hold across experiments and plant species. The buildup of meta-analytic results, however, allows us to go further and ask more general questions regarding selection pressures and plant responses, i.e. the types of questions on patterns of selection that are a major focus of modern evolutionary ecology. Key questions that can be explored are whether biotic or abiotic selective pressures are more important for plants, and if the strength of the response to those pressures depends on the plant trait studied.

Natural selection is an important force behind phenotypic differentiation across a wide range of plant traits (Kingsolver et al. 2001; Rieseberg et al. 2002). However, not all traits are expected to be targeted by selection in the same way. For example, traits closely related to fitness, such as life history traits, are expected to experience stronger selection than other types of traits (Merilä and Sheldon 1999). Tests of this idea have come to different conclusions, depending on the methodological approach. Kingsolver et al. (2001) compared selection gradients and differentials measured in wild populations across different types of traits, and found that morphological traits were subject to stronger selection than life history traits. Rieseberg et al. (2002) on the other

72 hand, compared the signature of selection with a more "historical" approach, using the

73 direction of effects of quantitative trait loci (the QTL sign test), and found evidence of

74 stronger and more consistent selection on life history than on morphological characters.

75   Contrasting results are not necessarily surprising because selective pressures are

76 expected to affect plant performance in complex ways (Bell 2010). For example, the

77 strength, form and direction of selection can vary in time (Grant and Grant 2002;

78 Siepielski et al. 2009; Kingsolver and Diamond 2011) and space (Linhart and Grant

79 1996; Schluter 2000; Herrera et al. 2006) but see also Morrisey and Hadfield (2012). In

80 addition, the type of selection pressure, whether biotic or abiotic, could also exert

81 different responses from plant traits. Biotic selective pressures depend on the

82 interactions with other organisms, such as predators or mutualists, whose distributions

83 and densities can vary rapidly and unpredictably and can therefore be expected to be

84 less consistent in strength, space and time (Linhart and Grant 1996; Thompson 2005).

85 Plant responses to biotic pressures could be then expected to be weaker and less

86 consistent across species and populations than to abiotic pressures. However, a recent

87 study suggests that biotically-selected traits are governed by fewer genes with a large

88 effect, which could allow populations to move faster among variable peaks in an

89 adaptive landscape (Louthan and Kay 2011). Although previous studies have tested for

90 differences between measures of selection on fitness traits compared to other types of

91 traits, to our knowledge no studies have specifically explored the potential differences

92 in selection when the pressures are biotic or abiotic.

93   We assembled here a database of diverse meta-analyses that allows us to

94 simultaneously test for the strength of the effects of biotic and abiotic selection

95 pressures and the responses of different types of traits (fitness versus non-fitness). We

96 also tested for the interaction between them, which could reveal differential effects of

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biotic or abiotic factors on different types of traits. We address these questions

quantitatively by performing a meta-analysis of published meta-analyses, or a second-

order meta-analysis, an approach that has been little used in ecology so far, but is

already common practice in the medical sciences (usually referred to as "umbrella

reviews" when various reviews are compared in narrative form, or "multiple treatment

meta-analysis" when multiple meta-analytic results are compared under specific

models; Caldwell et al. 2010; Ioannidis 2009). Specifically, we compare, with meta-

analytical techniques, a) the global effect sizes of meta-analyses of biotic versus abiotic

selection pressures, b) the global effect sizes of fitness versus non-fitness response

traits, and c) the interaction between them. Note that we are not dealing with data on

selection gradients or differentials (as defined by Lande and Arnold 1983), but with

studies that control or measure the selective factors and record their effect on plant

traits. In addition, we use our database to describe patterns in the published meta-

analysis literature on plant evolutionary ecology and detect potential biases towards

certain types of reviews.

Materials and methods

Clarification of the terminology used in the remaining of the article follows. The data

base used in our qualitative and quantitative analyses is composed of *meta-analyses*

mean effect sizes extracted from *publications* that may or may not include more than

one meta-analysis. Each meta-analysis in turn included *original case studies*. Data

points in our *second-order meta-analysis* are meta-analyses mean effect sizes and not

the original case studies. Methods are detailed below.

121   We compiled the data set of published meta-analyses on plant traits by performing

122   a literature search in the Web of Science with topic keywords "meta-analysis and plant"

123   (as of September 2011). We purged down the initial list of around 440 publications to

124   include meta-analyses that met the following requisites. a) Studies had to perform a

125   formal meta-analysis, that is, a comparison of weighted effect sizes across data sets. b)

126   Meta-analyses were revisions of the published literature designed to extract general

127   patterns. This excludes studies that used formal meta-analytical techniques to compare

128   various sets of original data. c) We excluded meta-analyses performed exclusively on

129   crop species under agricultural conditions, because a long history of artificial selection

130   might affect current response to selective pressures. d) We included only meta-analyses

131   focused on plant traits that can be measured in individuals. Community level (e.g.

132   species richness) or ecosystem level traits (e.g. litter decomposition) were not

133   considered. d) We also excluded allometric meta-analyses that were purely

134   morphological (e.g. trunk diameter vs. leaf area), when they had no clear evolutionary

135   implications.

136   We classified the remaining meta-analyses according to the type of response

137   variable studied (growth, physiology, reproduction or survival) and the type of selective

138   pressure examined (biotic or abiotic). Response variables were in turn grouped as

139   fitness variables (reproduction and survival) or non-fitness (physiological and growth

140   traits). This division might not seem straightforward, as it can be argued that growth or

141   development are fitness components as well. Our rationale follows that of Merilä and

142   Sheldon (1999), which assumes that reproductive traits and survival are more closely

143   related to fitness itself than other traits.

144   From each meta-analysis we extracted global effect sizes and their associated

145   sample sizes and sampling error variances to use them as weights. Sampling error

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146 variance is the square of the standard error, but these estimates are seldom reported in

147 the literature. Instead, 95% confidence interval of the effect size is usually provided and

148 half the width of the 95% CI divided by 1.96 is a good approximation to the standard

149 error. We did not include partial effect sizes (predictor factors) that subdivide data sets

150 already used to calculate a global effect (e.g. subdividing data sets to test the effect of

151 ant mutualisms on herbivory in shrubs versus herbs, Chamberlain and Holland 2009), to

152 avoid pseudoreplication. When several global effect sizes were provided by the same

153 publication to test separate response variable types (e.g. physiological, reproduction,

154 etc), we included all of them. For example, mutualism effects on growth and

155 reproduction of target plants were studied independently by Trager et al. (2010) and

156 therefore we included two global effects from this publication. Furthermore, if the

157 original meta-analysis mixed the types of response variables we were interested in, we

158 recalculated a global effect size for each variable type if the original data set was

159 available. For example, Bailey et al. (2009) reported effects of introgression on a

160 mixture of physiological, morphological, and reproductive response variables in

161 *Populus*. We recalculated global effect sizes for growth and physiological response

162 variables separately from their supplementary data set.

163   For our final second-order meta-analysis, we needed to transform individual meta-

164 analyses' effect sizes to a common metric. However, the most common effect metric

165 used in ecological studies, the log of the response ratio (lnRR) cannot be transformed

166 into other metrics in a straightforward way (M. Lejaunesse, pers. com.). We therefore

167 limited our quantitative analysis to meta-analyses reporting lnRR and closely related

168 metrics (e.g. percentage of change) and excluded those reporting metrics based on

169 standarized mean differences (i.e. Hedges d) or correlation coefficients. Because the

170 lnRR = $\ln(X^E) - \ln(X^C)$, i.e. the ratio of the outcome of an experimental group to that of

171   a control group, our database for the quantitative analysis is composed mostly of meta-

172   analyses of controlled experimental studies, but not exclusively, because some also

173   include original case studies using natural variation (e.g. Chamberlain and Holland

174   2009; Trager et al. 2010).

175   There was no significant correlation between effect size and sample size (r = 0.02,

176   df = 137, P= 0.81), suggesting against the biased publication of high effect sizes.

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178   Statistical analyses

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180   Because we were only interested in the strength of plant trait responses to selective

181   pressures, the sign of the effect sizes was not informative in our analysis. We therefore

182   used the absolute values of effect sizes (lnRR) to run Bayesian meta-analyses as

183   explained below. Using the absolute values could introduce an upward bias when

184   estimated effect sizes are non-significantly different from zero (Hereford et al. 2004).

185   However, we do not expect this to affect our comparisons, because around 80% of the

186   reported meta-analyses were significant. In addition, we are not testing for significance

187   in effect sizes, but rather for differences in their strength.

188   We first calculated an index of heterogeneity among meta-analyses (I^2; Higgins and

189   Thompson 2002) using the MCMCglmm R package as suggested by Nakagawa and

190   Santos (2012). Values of I^2 around 25%, 50% and 75% reflect small, medium and large

191   heterogeneity (Higgins et al 2003). For the second-order meta-analysis, we fitted

192   Generalized Linear Mixed Models using Markov chain Monte Carlo techniques with the

193   help of the MCMCglmm package for R (Hadfield 2010). The effect size was the

194   dependent variable in the model, and two types of weights were used: i) sample size and

195   ii) inverse of the sampling error variance. Both weighting strategies have been used in

196 social sciences (Hunter and Schmidt 2004 and references therein) as well as in ecology

197 (eg., van Groenigen et al 2011). Comparisons of the performance of both methods can

198 be found in Marín-Martínez and Sánchez-Meca (2010) and in Lajeunesse and Forbes

199 (2003).Weights passed to the *mev* argument of MCMCglmm (Hadfield and Nakagawa

200 2010). We ran 13000 MCMC iterations with a burn-in period of 3000 iterations and

201 convergence of the chain was tested by means of an autocorrelation statistic. The priors

202 used were nu=0 and V=I*1e+10, where I is an identity matrix of appropriate dimension.

203 The type of selective variable (biotic and abiotic) and the type of response variable

204 (fitness and non-fitness) were included as predictors in the MCMCglmm model,

205 including an interaction. Although separate global effect sizes could come from the

206 same publications, we decided against using the publication as a random grouping

207 factor in the model. This is because 1) separate meta-analyses reported in the same

208 publications are not necessarily non-independent, because they are derived from

209 different sets of original study cases, and 2) publications deal with only one of the

210 selective variable types (biotic or abiotic), so that including it as a random factor would

211 remove important variance from the main predictors unintentionally. The effect of

212 predictors was estimated by calculating the 95% credible interval of their posterior

213 distribution (Nakagawa and Cuthill 2007).

214

215 **Results**

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217 General patterns in the literature

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219 Our final data set included 196 meta-analyses based on more than 17800 original study

220 cases, reported in 51 publications (Table 1 and appendix). This sample reflects a bias in

221   the literature towards meta-analyses of non-fitness traits (154 versus 42 involving

222   fitness responses), and particularly towards those of non-fitness traits under abiotic

223   selection (102 studies). In contrast, only 9 meta-analyses in our data base dealt with

224   biotic characters under abiotic selective pressures.

225   Most abiotic selective pressures were climatic variables (111 vs. only 3 related to

226   disturbance). Among the climatic variables, there is a majority of meta-analyses dealing

227   with responses to elevated CO_2 (50 meta-analyses) and exposure to UV-B radiation

228   (25). Biotic pressures are all related to interactions, spanning from ant-plant

229   mutualisms (10 meta-analyses), to herbivory (19), interactions with plant neighbors

230   (11), and less often with plant-microbial interactions, pollinators, etc.

231   Finally, fitness responses are most often some measurement of reproductive output

232   (37 of 42 studies), while survival is the response variable in only 5 studies. In contrast,

233   within non-fitness variables there is a balance between growth and physiological

234   responses (77 each).

235

236   Quantitative analysis

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238   As explained above, we limit our quantitative analysis to the subset of meta-analyses in

239   our database reporting lnRR as the effect size (N= 139 meta-analyses in 30 publications.

240   Sampling error variance could only be obtained from 134 meta-analyses, see appendix).

241   We detected a large value of heterogeneity among meta-analyses (I^2 = 99.6%; [99.5,

242   99.7]), which justified using predictors. Results were very similar for both weighting –

243   sample size and variance- procedures. We found no significant interaction between the

244   type of selective variable (biotic and abiotic) and the type of response variable (fitness

245   and non-fitness) in their effect on effect sizes (posterior mean estimate = -0.010, 95%

246   CI [-0.251 to 0.208] for sample size weighted and -0.043 [-0.197, 0.079] for variance

247   weighted models). We therefore tested for the main effects of the two variables in a

248   model without interaction. It showed no significant differences in effect sizes between

249   fitness and non-fitness response variables (-0.059, [-0.172, 0.044] for sample size

250   weighted and -0.014 [-0.086, 0.051] for variance weighted models). However, there was

251   a significant effect of the type of selective variable analyzed, because biotic variables

252   elicit higher responses than abiotic ones (0.188 [0.104, 0.273] for sample size weighted

253   and 0.177 [0.120, 0.234]). Raw mean effect sizes and their standard errors are shown

254   in Fig. 1. These results are unchanged if we include response variables as physiology,

255   growth, or reproduction traits instead of grouping them as fitness or non-fitness.

Discussion

Our review of the meta-analytical literature of selection pressures on plants showed, on the one hand, that the majority of meta-analyses are studies of non-fitness traits and mostly on a few abiotic selection pressures such as increased CO_2 concentrations. On the other hand, these biases did not prevent a quantitative comparison of the effects of different selective pressures, which showed that the strength of responses to selection depends on the nature of selection (biotic versus abiotic factors) but, unexpectedly, not on the type of trait under selection. We discuss these results below.

Trends in the meta-analysis literature

Biases in our data base allowed us to detect biases in the meta-analysis literature. The majority of review studies are focused on non-fitness traits, i.e. on traits that are not directly related to reproduction or survival, and furthermore, on non-fitness traits under abiotic selection pressures. Certainly measuring a plant's reproductive output might be more difficult than measuring a morphological or physiological character and this can be one of the reasons for the unbalanced number of reviews. We suspect there is also a tradition of studying plant reproductive responses in a biotic context, and physiological and growth traits as influenced by abiotic environments (see Geber and Griffen 2003). These trends are reinforced by the recent boom of climate change studies, as reflected by the high number of CO_2 and UV radiation papers. The differential number of meta-analyses might then reflect a general bias in the plant literature. Louthan & Kay (2011), for example, also detected a bias towards abiotic-selected traits in a review of plant QTL mapping studies.

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282 Strength of biotic and abiotic selection on fitness and non-fitness traits

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284 Our approach to comparing the strength of selection on different types of traits differs

285 from other review papers (Kingsolver et al. 2001, Rieseberg et al. 2002, Geber and

286 Griffen 2003) in that we compare the results of multiple meta-analyses in a global,

287 second order meta-analysis that includes thousands of results published in the literature.

288 In addition, we do not focus on phenotypic selection as those articles, but on studies that

289 control biotic or abiotic environmental variation and measure the resulting fitness and

290 non-fitness responses. Because phenotypic selection studies do not formally measure

291 environmental variation, such studies would not be appropriate to test our hypothesis.

292 Still, we can compare our results on response variables to theirs. As opposed to those

293 previous findings, we did not detect differences in the strength of responses to selection

294 among different types of traits, either fitness or non-fitness. In contrast, when we looked

295 for differences in the responses to selection elicited by biotic versus abiotic traits, we

296 found a clear signal. Biotic-driven selection leads to stronger selection on traits in

297 general when compared to abiotic selection pressures, at least for plants. It is possible

298 that the biotic-abiotic comparison absorbs the differences between fitness and non-

299 fitness traits detected in previous studies, as both variables are collinear in our database

300 because of the biases described above.

301   The differential responses to biotic versus abiotic is a question that had been

302 basically unexplored. The main exception is the recent study by with Louthan and Kay

303 (2011), who compiled mapping studies on plant traits and compared the direction and

304 effect sizes of QTLs controlling biotic and abiotic-selected traits. Because they were not

305 dealing with selection studies directly, but rather with the consequences of selection on

306   the genetic architecture of traits, they classified traits *a-priori* as putatively abiotic- or

307   biotic-selected. Our study is the first that can confidently assign studies to the type of

308   selection pressure. Our reviewed studies report more immediate responses and the

309   results are therefore less historical than a QTL comparison, but in spite of the difference

310   in approaches, the two studies found consistent results. Louthan and Kay (2011) found

311   QTL's of larger effect associated with biotic-selected traits, and we found stronger

312   observed responses of traits under biotic pressures. Both results are expected for traits

313   that are under variable selective pressures, as can be the case for biotic selective agents.

314   Biotic agents and interactions can vary strongly in space and time (Thompson 2005 and

315   references therein), and consequently produce complex selective landscapes with

316   multiple peaks or peaks that in turn shift in time and space. Such selective scenario can

317   produce phenotypic responses that are stronger than under more subtle abiotic changes,

318   and in turn select for QTLs of major effects.

319   To further explore the relative importance of biotic versus abiotic factors on plant

320   character evolution it is clear that a higher diversity of studies would be very useful. In

321   particular, case studies and meta-analyses in the under-represented categories (fitness

322   traits under abiotic selection and non-fitness traits under biotic selection) would be very

323   valuable. In addition, fully factorial case studies on the effects of biotic and abiotic

324   pressures on both fitness and non-fitness traits in individual species are scant but

325   potentially very informative.

326

327   Guide for future meta-analyses of meta-analysis

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329   The broad use of formal meta-analytical techniques in plant ecology has undoubtedly

330   contributed to our capacity for summarizing and extracting general results, based on the

331   strength of combining many varied individual studies. We here take the next step of

332   combining effect sizes of meta-analyses on diverse plant systems and traits in a second

333   order meta-analysis. This approach is already frequently used in the health sciences,

334   particularly to answer clinical questions, where for example different treatments for the

335   same disease need to be compared but results are reported in independent reviews

336   (Ioannidis 2009; Becker and Oxman 2011). Multiple-treatment meta-analysis is used to

337   formally compare meta-analytic results in a network approach that incorporates direct

338   and indirect comparisons of clinical treatments (Hasselblad, 1998; Caldwell et al. 2010).

339   Our analysis is a simplified version of such models.

340   One advantage of the approach of meta-analyzing meta-analyses is that it allows a

341   high level of generalization using a very large number of individual case results already

342   summarized in meta-analyses (in our case, more than 17800) that would be very

343   impractical to attempt with the original studies. Most meta-analyses, except perhaps the

344   most recent ones, do not list each individual study case included and their associated

345   effect size, sample size and variance, all required for a new meta-analysis based on the

346   original studies. In a recent article that used published meta-analyses to find groups of

347   papers on specific topics and extract individual study information (Barto and Rillig

348   2011), the authors report that they had to limit their analysis to a small fraction of the

349   available publications, because few report the necessary data for each case study. In our

350   case, using the original data would then imply going back to each case study and

351   repeating the work done by meta-analytical studies. Another advantage of the second-

352   order approach of using published meta-analyses compared to searching for original

353   case studies is that meta-analyses are prepared by expert authors, who identify the

354   relevant questions on each topic and the appropriate case studies to answer them. In a

355 broad second-order meta-analysis like ours, such level of expertise is left to the original

356 reviews.

357   Nonetheless, some aspects need to be considered carefully before combining

358 review studies in second-order meta-analysis. First, it is possible that the same

359 individual original studies are included in more than one of the meta-analytic

360 publications available on a given topic. Our questions here were so broad and the

361 number of individual studies on different topics so large, that it is unlikely that this form

362 of pseudoreplication has affected our conclusions. Smaller and more focused meta-

363 meta-analyses should probably be more concerned with excluding replicated results.

364 Care should also be taken not to include meta-analyses that were not careful about

365 another possible form of pseudoreplication, i.e, using the same case studies (and same

366 experimental individuals) to conclude on different effects. Second, there are statistical

367 problems with the conversion of effect sizes to a single common metric, as explained in

368 the Methods section. This can be a problem in ecological studies particularly, because a

369 variety of effect sizes are commonly used and in particular response ratios, whose

370 statistical properties have not been fully studied yet. Because of this problem, we had to

371 limit our quantitative analysis to a single family of effect size metrics and exclude many

372 potentially informative meta-analyses. Finally, future second-order meta-analyses

373 addressing evolutionary issues should ideally include phylogenetic-informed effect

374 sizes (Verdú and Traveset 2004), because of the ubiquity of phylogenetic signal in

375 ecological traits (Blomberg et al. 2003).

376

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474   **Figure legends**

475

476   Figure 1. Raw mean and 1 standard error of effect sizes (lnRR) for meta-analyses

477   classified according to the type of selective variable (biotic and abiotic) and the type of

478   response variable (fitness and non-fitness). Sample sizes for each group are included.

479   Table 1. Number of meta-analyses in each category of selective pressures and trait

480   response types included in this revision. Details and references are in appendix 1.

481

Selective pressure	Response trait type	Meta-analyses in this study
Biotic	Fitness	
	reproduction	28
	survival	5
	Non-fitness	
	physiology	14
	growth	38
Abiotic	Fitness	
	reproduction	9
	survival	0
	Non-fitness	
	physiology	63
	growth	39

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