

# FULL TITLE: Incorporating Truncating Variants in *PALB2*, *CHEK2* and *ATM* into the BOADICEA Breast Cancer Risk Model

## SHORT TITLE: Rare Variants in the BOADICEA Breast Cancer Risk Model

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## **CONFLICT OF INTEREST**

There are no known conflicts of interest.

## ABSTRACT

### *Purpose:*

The proliferation of gene-panel testing precipitates the need for a breast cancer (BC) risk model that incorporates the effects of mutations in several genes and family history (FH). We extended the BOADICEA model to incorporate the effects of truncating variants in *PALB2*, *CHEK2* and *ATM*.

### *Methods:*

The BC incidence was modelled via the explicit effects of truncating variants in *BRCA1/2*, *PALB2*, *CHEK2* and *ATM* and other unobserved genetic effects using segregation analysis methods.

### *Results:*

The predicted average BC risk by age 80 for an *ATM* mutation carrier is 28%, 30% for *CHEK2*, 50% for *PALB2*, 74% for *BRCA1* and *BRCA2*. However, the BC risks are predicted to increase with FH-burden. In families with mutations, predicted risks for mutation-negative members depend on both FH and the specific mutation. The reduction in BC risk after negative predictive-testing is greatest when a *BRCA1* mutation is identified in the family, but for women whose relatives carry a *CHEK2* or *ATM* mutation, the risks decrease slightly.

### *Conclusions:*

The model **may** be a valuable tool for counselling women who have undergone gene-panel testing for providing consistent risks and harmonizing their clinical management. A web-application can be used to obtain BC- risks in clinical practice (<http://ccge.medschl.cam.ac.uk/boadicea/>).

Keywords: breast cancer; risk prediction; BOADICEA; BRCA1; BRCA2; PALB2; CHEK2; ATM; user interface; gene-panel

## INTRODUCTION

Breast cancer exhibits strong familial aggregation, such that the risk of the disease increases with increasing number of affected relatives. First degree relatives of women diagnosed with breast cancer are at approximately 2 times greater risk of developing breast cancer themselves than women from the general population<sup>1</sup>. Many breast cancer susceptibility variants have been identified to date. Approximately 15-20% of this excess familial risk is explained by rare, high penetrance mutations in *BRCA1* and *BRCA2*<sup>2,3</sup>. Other rare, intermediate risk variants (e.g. mutations in *PALB2*, *CHEK2* and *ATM*) are estimated to account for ~5% of the breast cancer familial aggregation<sup>4-6</sup>, and the common, low-risk alleles identified through genome-wide association studies to account for a further 14% of the familial risk<sup>7,8</sup>.

To provide comprehensive genetic counselling for breast cancer, it is important to have risk prediction models that take into account the effects of all the known breast cancer susceptibility variants, and also account for the residual familial aggregation of the disease. Some existing genetic risk prediction algorithms incorporate the explicit effects of *BRCA1* and *BRCA2* mutations, including BRCAPRO<sup>9</sup>, IBIS<sup>10</sup> and BOADICEA<sup>3,11</sup>. BOADICEA accounts for the residual familial aggregation of breast cancer in terms of a polygenic component that models the multiplicative effects of a large number of variants each making a small contribution to the familial risk<sup>3</sup>.

Next generation sequencing technologies that enable the simultaneous sequencing of multiple genes through gene-panels<sup>12,13</sup> have now entered clinical practice. However, the clinical utility of results from such genetic testing remains limited as none of the currently available risk prediction models incorporate the simultaneous effects of the rare-intermediate risk variants and other breast cancer risk factors, in particular explicit family history. As a result, providing risk estimates for women who carry these mutations, and their relatives, is problematic<sup>6</sup>.

In this paper, we describe an extension to the BOADICEA model to incorporate the effects of intermediate risk variants for breast cancer, specifically loss of function mutations in the three genes

for which the evidence for association is clearest and the risk estimates most precise: *PALB2*, *CHEK2* and *ATM*. The resulting model allows for consistent breast cancer risk prediction in unaffected women on the basis of their genetic testing results and their family history.

## MATERIALS AND METHODS

### Breast Cancer Incidence in BOADICEA

We build on the existing BOADICEA model<sup>2,3,11</sup>. Briefly, in this model, the breast cancer incidence,  $\lambda_i(t)$ , for individual  $i$  at age  $t$  is assumed to be birth cohort specific, and to depend on the underlying *BRCA1* and *BRCA2* genotypes and the polygenotype through a model of the form:

$$\lambda_i(t) = \lambda_0(t) \exp(\beta_1(t)G_{1i} + \beta_2(t)(1 - G_{1i})G_{2i} + P_i(t)), \quad \text{Equation(1)}$$

where  $\lambda_0(t)$  is the baseline incidence for the cohort,  $G_{1i}$  is an indicator variable taking values 1 if a *BRCA1* mutation is present and 0 otherwise, and similarly  $G_{2i}$  for *BRCA2*.  $\beta_1(t)$  and  $\beta_2(t)$  represent the age-specific log-relative risks associated with *BRCA1* and *BRCA2* mutations respectively, relative to the baseline incidence (applicable to a non-mutation carrier with a zero polygenic component) and where  $P_i(t)$  is the polygenic effect, assumed to be normally distributed with mean 0 and variance  $\sigma_p^2(t)$ .

The effects of mutations in *BRCA1* and *BRCA2* are modelled *through* a single locus “major gene” with three alleles (*BRCA1*, *BRCA2* and wild-type). The *BRCA1* and *BRCA2* alleles are assumed to be dominantly inherited<sup>14</sup>. As a further simplification, carriers of both the *BRCA1* and *BRCA2* alleles are assumed to be susceptible to *BRCA1* risks. These simplifications reduce the number of possible “major” genotypes from 9 to 3: (1) being a *BRCA1* mutation carrier; (2) being a *BRCA2* mutation carrier; and (3) being a non-mutation carrier.

The BOADICEA genetic model uses the Elston-Stewart peeling algorithm to compute the pedigree likelihood<sup>15,16</sup>. As a result, the number of computations increases exponentially with the number of possible genotypes in the model. To maintain computational efficiency, we incorporated the effects of risk-conferring variants in *PALB2*, *CHEK2* and *ATM* into the model by introducing an additional allele for each gene (representing a mutation in that gene) to the *BRCA1/2* major gene locus, resulting in a locus with 6 alleles. In comparison with a model that has a single locus for each gene, this approximation can be justified by the low allele mutation frequencies for all genes (Table 1), because the probability of carrying mutations in more than one of these genes is low<sup>17</sup>, relative to the probability of carrying one or no mutation. Currently, few published data describe the cancer risks to individuals carrying more than one mutation<sup>18</sup>. Here we have assumed that the risks follow a dominant model, with the order of precedence being *BRCA1*, *BRCA2*, *PALB2*, *CHEK2*, *ATM* and wild-type. Under this model, in the presence of a mutation in one gene, no additional risk is conferred by a second mutation in another gene lower in the dominance chain.

## Relative Risks for Female Breast Cancer

We extended the model for the breast cancer incidence to incorporate the effects of rare variants in *PALB2*, *CHEK2* and *ATM*, such that:

$$\lambda_i(t) = \lambda_0(t) \exp \left( \sum_{\alpha=1}^5 \beta_{\alpha}(t) \prod_{\gamma=1}^{\alpha-1} (1 - G_{\gamma i}(t)) G_{\alpha i}(t) + P_{Ri}(t) \right), \quad \text{Equation(2)}$$

where  $\lambda_0$ ,  $\beta_1(t)$ ,  $G_{1i}$ ,  $\beta_2(t)$  and  $G_{2i}$  are as described in Equation(1), and  $G_{3i}$ ,  $G_{4i}$ , and  $G_{5i}$ , are indicator variables taking values 1 if a mutation is present and 0 otherwise, for *PALB2*, *CHEK2* and *ATM* respectively.  $\beta_3(t)$ ,  $\beta_4(t)$  and  $\beta_5(t)$  represent the age-specific log-relative risks associated with *PALB2*, *CHEK2* and *ATM* mutations respectively, relative to the baseline incidence (applicable to a non-mutation carrier with a zero polygenic component).  $P_{Ri}(t)$  is the residual polygenic component, with mean 0, and variance  $\sigma_R^2(t)$ , explained below.

To implement the model, we assumed the mutation frequencies and relative risks (RRs) summarised in Table 1. The RR estimates (Table 1) represent estimates relative to the population incidences and are therefore RR over all polygenic effects. Multiplying the RRs (Table 1) by the cohort and age specific incidences yields the average incidences in carriers of *PALB2*, *CHEK2* and *ATM* mutations over all polygenic effects. To obtain  $\beta_3(t)$ ,  $\beta_4(t)$  and  $\beta_5(t)$ , we constrained the overall incidences (using as weights the major genotype and polygenic frequencies) to agree with the population breast cancer incidence, for each birth cohort separately. This process is described in detail elsewhere<sup>14</sup>

To ensure that the familial risks predicted by this extended model remain consistent with the previous model, we adjusted the variance of the polygenic component to account for the fact that the contributions of *PALB2*, *CHEK2* and *ATM* to the genetic variance are now explicitly accounted for in the major gene, following the process described in<sup>19</sup>. Briefly, the total polygenic variance ( $\sigma_p^2$ ) was decomposed into the sum of the known variance ( $\sigma_K^2$ ), due to the three variants, and residual variance ( $\sigma_R^2$ ),

$$\sigma_p^2 = \sigma_K^2 + \sigma_R^2,$$

The known variance,  $\sigma_K^2$ , can be calculated as a function of the joint RRs and mutation frequencies of the three risk variants<sup>19</sup>. This assumes that the effect of each variant and the effect of residual polygene (which describes residual familial aggregation) are multiplicative. This is in line with recent findings for *PALB2* mutations<sup>20</sup>. This model is also consistent with the higher RR for *CHEK2* 1100delC for breast cancer based on familial cases<sup>21,22</sup>, the higher RR for bilateral breast cancer<sup>23</sup>, and the increased risk of breast cancer in relatives of breast cancer patients who are *CHEK2* carriers<sup>24</sup>. A higher RR for familial breast cancer for *ATM* carriers has also been found, though the data are more limited for *ATM*<sup>25</sup>.



## PALB2 Characteristics

As there are very limited data from case-control studies for *PALB2* or *ATM*, we used alternative family-based data. Age dependent RRs of female breast cancer for carriers of loss of function variants in *PALB2* were taken from a large collaborative family-based study<sup>20</sup>. With the exception of specific Nordic founder mutations, data on mutation frequencies in the general population are sparse. We assumed a mutation allele frequency of 0.057% (or a mutation frequency of ~0.11% (10/8705)) based on data from targeted sequencing of 8705 controls from the UK (unpublished data). This is close to the average estimate across published estimates<sup>26</sup>.

## CHEK2 Characteristics

Most existing data describe the *CHEK2* 1100delC variant, which is the most common truncating variant in northern European populations<sup>27</sup>. *CHEK2* 1100delC has been evaluated in many case-control studies<sup>22,28</sup>. As a result, we based the *CHEK2* estimates on the *CHEK2* 1100delC carrier estimates from a meta-analysis<sup>28</sup>. We assumed that the allele frequency of the 1100delC mutations was 0.26%, the combined frequency across unselected population controls of European ancestry<sup>28</sup>. The missense variant *CHEK2* I157T has also been shown to confer increased risks of breast cancer<sup>7,29</sup>. However, these risks are much lower than those conferred by *CHEK2* 1100delC, and are closer to those conferred by common susceptibility variants. Consequently, *CHEK2* I157T has been included in published Polygenic Risk Scores with other common-low penetrance alleles, and it will be included in future model extensions<sup>30</sup>.

There is some evidence that the RRs for breast cancer in *CHEK2* 1100delC carriers decline with age<sup>22</sup>. However, since age-specific estimates are currently imprecise, we used a single RR estimate across all ages.

## ATM Characteristics

We obtained estimates for truncating mutations in *ATM* from a combined analysis of three estimates from cohort studies of relatives of Ataxia-Telangiectasia (A-T) patients (Table 1)<sup>6</sup>. The large majority of A-T patients carry two truncating *ATM* mutations, and relatives of A-T patients are therefore known to have a high probability of being carriers of an *ATM* mutation. We assumed that the allele frequency of truncating variants in *ATM* was 0.19% based on data from UK controls<sup>25</sup>. As for *CHEK2*, there is some evidence of a decline in RR with age<sup>31</sup>, but in the absence of relative age-specific estimates, we used a single estimate across all ages.

## Relative Risks for Other Cancers

In addition to the risks of female breast cancer, BOADICEA takes into account the associations of *BRCA1* and *BRCA2* mutations with the risks of male breast cancer, ovarian cancer, male and female pancreatic cancer, and prostate cancer<sup>3</sup>. Several studies have investigated the associations of the truncating variants in *PALB2*, *ATM* and *CHEK2* 1100delC with the risks of these cancers (and others)<sup>20,29,31,32</sup>. However, none of the studies have provided convincing evidence of association for any of these cancers, and accurate penetrance estimates are currently lacking for those cancers that may have associations. Therefore, for the purpose of the current implementation, we assumed that these mutations are not associated with elevated risks of other cancers.

## Incorporating Breast Tumour Pathology Characteristics

Previous studies<sup>2,11</sup> describe the incorporation into BOADICEA of differences in tumour pathology subtypes between *BRCA1*, *BRCA2* and non-carrier breast cancers. Specifically, BOADICEA includes information on tumour oestrogen receptor (ER) status, triple negative (oestrogen, progesterone and HER2 negative) (TN) status, and the expression of the basal cytokeratin markers CK5/6 and CK14.

Breast cancers in *CHEK2* 1100delC mutation carriers have been found to be ER-positive, at a greater proportion compared to tumours in non-*CHEK2* mutation carriers (ER-positive tumours in *CHEK2*

1100delC carriers=88% (290/331) vs 78% in general population)<sup>33</sup>. Reliable data pertaining to the TN and basal cytokeratin receptor status are not currently available. Therefore, in the current implementation, we only incorporated differences by breast cancer ER-status for *CHEK2* 1100delC carriers, assuming that 88% of tumours in carriers are ER-positive. Age specific distributions were not available.

Published data on the prevalence of these tumour subtypes in *PALB2* associated breast cancers are currently sparse, and although some differences compared to the general population have been reported, these are based on small numbers<sup>20,34</sup>. Currently there are no available data pertaining to tumour pathology subtype distributions for carriers of *ATM* truncating mutations. We therefore assumed that the tumour subtype distributions for *PALB2* and *ATM* mutation carriers are the same as in the general population.

### **Mutation Screening Sensitivity**

We have introduced separate mutation test screening sensitivities for *PALB2*, *CHEK2* and *ATM* to allow for the fact that some risk-conferring variants in these genes may be missed by current screening methods. In the BOADICEA Web Application (BWA), we assumed default values of 90% for *PALB2* and *ATM* truncating variants, and 100% for the *CHEK2* 1100delC variant. However, these values can be customised by users, where appropriate. The specificity of mutation testing was assumed to be 100%.

## **RESULTS**

Fig 1(a) (and Fig S1) shows the implied average cumulative breast cancer risks predicted by BOADICEA by mutation status, on the basis of the assumed RR parameters for an unaffected female aged 20 born in 1975. The predicted average breast risk by age 80 for a *CHEK2* mutation carrier was 29.9%, 28.2% for *ATM*, 50.1% for *PALB2*, 73.5% for *BRCA1* and 73.8% for *BRCA2*.

On the basis of the assumed mutation frequencies and RRs and modelling assumptions, the known polygenic variance ( $\sigma_K^2(t)$ ) due to the effects of *PALB2*, *CHEK2* and *ATM* are given in Table 2. The age dependence of the variances due to *PALB2*, *CHEK2* and *ATM* is a consequence of the fact that relative risks vary with age (in particular for *PALB2*) and the age dependence of the frequency of mutation carriers among the unaffected population, which decreases with age (elimination effect). The proportion of polygenic variance accounted for by the three genes varied from 3.0% at age 25 to 9.8% at age 75.

### **Mutation Carrier Probabilities**

Fig 2 shows the mutation carrier probabilities predicted by BOADICEA for each gene, for (a) a female with unknown family history as a function of her age of cancer diagnosis, and (b) for a 30 year old female diagnosed with breast cancer, whose mother has had breast cancer, as a function of her mother's age at diagnosis (also given in Tables S1 and S2). The mutation carrier probabilities for *ATM* and *CHEK2* did not show a marked change with age at diagnosis (reflecting the assumption of a constant RR by age), but the mutation carrier probabilities decreased with age for *PALB2*, though less markedly than for *BRCA1* or *BRCA2*. As expected, the mutation carrier probabilities were higher for women with a family history, but the effect was more marked for *BRCA1*, *BRCA2* and *PALB2* than for *CHEK2* or *ATM*.

### **Predicted Cancer Risks for Mutation Carriers are Family History Specific**

In our model, the residual polygenic component was assumed to act multiplicatively with *PALB2*, *CHEK2* and *ATM* mutations on the risk of developing breast cancer. As a result, the breast cancer risks for mutation carriers will vary by family history of breast cancer. Fig 1 shows the predicted cumulative breast cancer risk for a 20 year old UK woman born in 1975 by her mutation status. In (a) the woman was assumed to have unknown family history; in (b) to have a mother affected with breast cancer at age 40; and in (c) to have a mother and sister who are cancer free at ages 70 and 50 respectively. The figures show clearly that the predicted breast cancer risks increased with

increasing number of affected relatives, and depend on the phenotypes of the unaffected family members. For example, although the average breast cancer risks by age 80 for *CHEK2* and *ATM* mutation carriers were lower than 30% (a common criterion for “high” risk, used for example in the NICE guideline<sup>35</sup>) the breast cancer risk exceeded this threshold when a mutation carrier had a family history of breast cancer (e.g. 42.6% for an *ATM* and 44.7% for *CHEK2* mutation carrier with an affected mother). Comparing Figures 1 (c) and 1 (a) we see that the risk for a woman with no history of breast cancer is lower than the average breast cancer risk.

### The Effect of Negative Predictive Testing

The extended BOADICEA model can also be used to calculate the predicted risks in families in which mutations are identified, but other family members test negative for the mutation. This is demonstrated for a number of family history scenarios in Fig 3, which each depend on the mutation status of the proband and her mother. The predicted risks for mutation negative family members depend on both the family history of breast cancer and the specific mutation identified in the family. Thus for families with a history of breast cancer, namely (c), (e) and (g), the reduction in breast cancer risk after negative predictive testing is greatest when a *BRCA1* mutation was identified in the family, with the risks being close to (though still somewhat greater than) the correspond risks based on population rates. This effect was most noticeable for women with a strong family history. The reduction in risk in women whose mother carried a *BRCA2* or *PALB2* mutation is less marked, while for women whose mother carried a *CHEK2* or an *ATM* mutation, the risks decreased only slightly with a negative predictive test, even for women with a strong family history. For a woman with no history of breast cancer(Figure 3 (a)), her risk on the basis of family history alone (i.e. in an untested family) was slightly lower than the population risk. After negative predictive testing her predicted risk decreased further. The biggest decrease was observed when a *BRCA1* mutation was identified in the mother.

## Updates to the BOADICEA Web Application

We have now updated the BWA (<http://ccge.medschl.cam.ac.uk/boadicea/>) to accommodate the extensions to the BOADICEA model presented here. The BWA enables users to either build a pedigree online for processing, or to upload a text file containing one or more pedigrees for processing. When users build an input pedigree online, the program now enables users to specify *PALB2*, *CHEK2* and *ATM* genetic test results in risk calculations. Similarly, we have extended the BOADICEA import/export format (described in Appendix A of the BWA v4 user guide: [https://pluto.srl.cam.ac.uk/bd4/v4/docs/BWA\\_v4\\_user\\_guide.pdf](https://pluto.srl.cam.ac.uk/bd4/v4/docs/BWA_v4_user_guide.pdf)) so that users can include this information in the text files that they upload for processing.

## DISCUSSION

Cost-effective sequencing technologies have brought multi-gene panel testing into mainstream clinical care<sup>6,13,36</sup>. Although several established breast cancer susceptibility genes are included in these panels, their clinical utility is limited by the lack of risk prediction models that consider the effects of mutations in these genes and other risk factors for breast cancer, in particular family history. Here, we present an extended BOADICEA model that incorporates the effects of rare protein truncating variants in *PALB2*, *CHEK2* and *ATM*. This is the first breast cancer risk prediction model to include the explicit effects of breast cancer susceptibility genes other than *BRCA1* and *BRCA2*, and it can be used to provide comprehensive risk counselling on the basis of family history and mutation screening in the five genes. The model can also be used to predict future risks of developing breast cancer and the likelihood of carrying truncating mutations in any of the five genes.

The extended BOADICEA model is based on a number of assumptions. To ensure the model is computationally efficient we used a single "major" locus with six alleles representing the truncating variants in the five genes and a wild-type allele. In comparison with a genetic model consisting of 5 separate loci each with 2 alleles, this should be a reasonable approximation as all the variants are

rare. Algebraically leading errors should be proportional to the product of any two allele frequencies, which is of the order of  $10^{-6}$ . Our empirical measurements confirm that this is the case, with leading differences in the 6<sup>th</sup> significant figure. However, it is possible the errors will be greater in families segregating more than one rare variant. It also represents a substantial reduction in the number of genotypes (36 V's 1024), and hence in execution time; we measure execution time to be reduced by a factor of 21000. These simplifications will become more critical as the number of susceptibility genes included in the model increases. A previous study<sup>6</sup> identified six other genes for which the association with breast cancer was well established (*TP53*, *PTEN*, *STK11*, *CDH1*, *NF1* and *NBN*), and this list is likely to increase as large exome- and targeted- sequencing studies are completed.

In the absence of robust data on the risks conferred by carriers of 2 or more truncating variants (in different genes), we assumed that dual mutation carriers develop breast cancer according to incidences for the higher penetrance gene. Recent evidence suggests that gene-gene interaction between *CHEK2*, *ATM*, *BRCA1* and *BRCA2* may not be multiplicative (indeed a multiplicative model would clearly be implausible for *BRCA1* and *BRCA2*, since it would predict an extremely high risk to double mutation carriers at very young ages)<sup>18</sup>. This may reflect the biological relationships between the proteins encoded by the genes. The proteins encoded by all five genes play roles in DNA repair, and loss of function mutations in these genes are predicted to impair DNA repair. Our implementation would be consistent with a model where if the pathway is disrupted by one mutation, further disruption by a lower penetrance mutation would not lead to an increased risk.

Although there is strong evidence that mutations in *PALB2*, *CHEK2* and *ATM* confer increased risk of breast cancer in females<sup>6</sup>, there are currently no precise risk estimates for the other cancers considered by BOADICEA (male breast, ovarian, pancreatic or prostate), or indeed other cancers. However, several studies have provided tentative evidence of associations<sup>20,32</sup>. Due to the lack of precise cancer risk estimates, we have assumed no association between truncating variants in

*PALB2*, *CHEK2* and *ATM* (i.e. RR=1). If there are true associations between the *PALB2*, *CHEK2* and *ATM* truncating variants and other cancer risks, we expect that *PALB2*, *CHEK2* and *ATM* mutation carrier probabilities may potentially be underestimated in families where other cancers occur. However, our implementation is flexible enough that should accurate risk estimates for other cancers become available, they can easily be included.

BOADICEA allows data on cancer tumour characteristics to be taken into account, as we have done previously for *BRCA1* and *BRCA2*<sup>2,11,37</sup>. The provision of subtype-specific risks can be useful for genetic counselling and may guide chemoprevention. However, data on the additional genes are currently sparse. In this model, we incorporated a higher probability of ER-positive tumours in *CHEK2* 1100delC carriers, relative to non-carriers<sup>33</sup>. Some studies have suggested differences in the characteristics of tumours from *PALB2* mutation carriers and non-carriers, but larger studies will be required to establish such differences<sup>20,34</sup>.

In this model, we considered only the effects of truncating variants in *PALB2*, *ATM* and of the *CHEK2* 1100delC variant, for which reasonably robust breast cancer risk estimates are available. In doing this, we are making the usual simplification that all truncating variants in these genes confer similar risks. While there is no evidence to contradict this, it may change as further data accumulate. In addition, there is evidence that missense variants in both *CHEK2* and *ATM* also confer elevated breast cancer risks, but that the risks that they confer can differ from the risks associated with truncating variants. For example, the *ATM* c.7271T>G missense variant has been reported to confer a higher risk than truncating variants, but the confidence intervals associated with this estimate are currently wide<sup>38</sup>. It has been suggested that other rare, evolutionarily unlikely missense variants in *ATM* are also associated with increased breast cancer risks<sup>39</sup>. Future extensions of BOADICEA can accommodate such differences on the basis of more precise cancer risk estimates. In *CHEK2*, the missense variant Ile157Thr has been associated with a lower risk than the 1100delC variant<sup>40</sup>. This variant has been recently incorporated into a polygenic risk score on the basis of common genetic



variants<sup>30</sup> and we expect to incorporate this into BOADICEA through ongoing extensions that will include the effects of common genetic variants. The model could also be applicable for other truncating variants in *CHEK2*, under the assumption that they confer similar risks to the 1100delC variant. However, the available data are scarce and some modification of the mutation frequencies may be required.

Under the BOADICEA model, women testing negative for known familial mutations (true negatives) and who have family history of breast cancer are predicted to be at higher risk of breast cancer than the general population. The level of risk depends on both family history and the specific mutation identified in the family. So far, epidemiological studies have reported estimates for “true negatives” only in the context of families with *BRCA1* and *BRCA2* mutations<sup>41-46</sup> but the estimated relative risks (compared to the population risks) have been found to vary widely. Moreover, all the reported estimates are associated with wide confidence intervals because the studies have been based on small sample sizes. The reported estimates are summarised in Table S3. To provide a direct comparison with the predicted risks by BOADICEA we have also included the implied relative risks for the “true negative” women in Figure 3 relative to the population risk in Table S3. These are all in line with the published estimates for true negatives. Therefore, the predictions by BOADICEA are consistent with published epidemiological data. It is worth noting that if the true relative risks for the “true negatives” in families with *BRCA1* and *BRCA2* mutations are in line with those predicted by BOADICEA, very large prospective studies of “true negatives” will be necessary to demonstrate significant associations.

The current model is a synthetic model, based on segregation analyses of families in the UK together with risk estimates derived from studies of European populations. We have previously implemented procedures for extrapolating the model to populations with different baseline incidence rates, on the assumption that the RRs conferred by the genetic variants in the model are independent of the

population<sup>11</sup>. Thus, the model should be broadly applicable to developed populations of European ancestry, but its applicability to populations with lower incidence rates, and populations of non-European ancestry, has yet to be evaluated. The implementation also allows the allele frequencies to be adjusted. This may be particularly relevant for *CHEK2*; in European populations the founder 1100delC variant accounts for the majority of carriers of truncating variants, and its frequency varies widely across populations.

The extended BOADICEA model presented here has addressed a major gap in breast cancer risk prediction, by including the effects of truncating variants in *PALB2*, *CHEK2* and *ATM* that are included in widely used commercial gene-panels. The model **could** be a valuable tool in the counselling process of women who have undergone gene-panel testing for providing consistent breast cancer risks and thus harmonizing the clinical management of at risk individuals. Future studies should aim to validate this model in large prospective cohorts with mutation screening information and to evaluate the impact of the risk predictions on decision making.

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## TABLE AND FIGURE LEGENDS:

**Table 1.** Mutation frequency and relative risks (RR) for loss of function variants in *PALB2*, *CHEK2* and *ATM*. The RRs for *PALB2* are taken from<sup>20</sup>. The mutation frequency for *PALB2* is taken from a private communication from Easton and Pharaoh based on data from unaffected individuals from the UK. Relative risks for *CHEK2* and *ATM* are taken from<sup>6</sup>. The allele frequency for *CHEK2* is taken from<sup>28</sup>, and the allele frequency for *ATM* is taken from<sup>31</sup>.

**Table 2.** The variance explained by *PALB2*, *CHEK2* and *ATM* and the percentage of the overall polygenic variance explained by all three combined.

**Figure 1. BOADICEA Breast Cancer Risk by Mutation Status and Family History.** BOADICEA risk by mutation status for a female in the UK age 20 born in 1975: (a) with unknown family history (i.e. for the average female in the population); (b) with her mother affected at age 40; (c) with her mother and sister unaffected at ages 70 and 50 respectively. No testing assumed in other family members, in all cases.

**Figure 2. BOADICEA Mutation Carrier Probabilities.** BOADICEA mutation carrier probabilities for a female in the UK, born in 1975: (a) with unknown family history as a function of her breast cancer diagnosis age; (b) who was diagnosed with breast cancer at age 30 and whose mother was diagnosed with breast cancer, as a function of her mother's age at diagnosis.

**Figure 3. BOADICEA Breast Cancer Risk for Negative Testing by Family History.** The predicted risk of breast cancer for a 20 year old female in the UK, born in 1975 by her mother's mutation status, for different family histories. The predicted risk is shown for four different family histories. The graphs

on the right hand side correspond to the pedigrees on the left hand side. The figures show the predicted risks for a proband (shown with an arrow) in families without any mutation testing in the five genes i.e. this corresponds to the predicted risk on the basis of family history information alone (grey curves). The rest of the curves correspond to the cases where the proband is assumed to be negative for the mutation identified in the family. To enable direct comparisons, the proband is assumed to be 20 years old in all examples.