

## [The relationship between exercise-induced and non-stimulated 24-hour growth hormone release](#)

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### Abstract:

The purpose of this study was to quantify the relationship between the acute exercise-induced growth hormone (GH) response, gender, fitness, age, and body composition with non-exercise stimulated (NES) 24-hr integrated GH concentration (IGHC). Twenty-nine subjects (16 males, 13 females) completed a 24-hr control session without exercise (NES) on one occasion and a 30-min exercise session on a separate occasion. The GH release was analyzed from 10-min intervals using trapezoidal integration during both sessions. Multiple regression analysis in males revealed that significant variability in NES 24-hr IGHC could be adequately explained by subject age and BMI together ( $P<0.05$ ), but could not be explained by exercise-induced peak GH or 30-min of exercise-induced IGHC on an unrelated day ( $P=0.750$ ). However, significant variability in NES 24-hr IGHC in females could be uniquely associated with subject age ( $P<0.05$ ) as well as peak GH and 30-min of exercise-induced IGHC ( $P<0.05$ ). The findings indicate that in the females only the constant load exercise-induced growth hormone response can adequately predict total 24-hr growth hormone output on a separate day without an exercise stimulus. The data further exemplify gender disparities in exercise-induced growth hormone release and the importance of exercise for females in regards to enhancing total 24-hr growth hormone concentrations.

**Keywords:** Body Composition | Sex | Hormones | Obesity

### Article:

## INTRODUCTION

Growth hormone (GH) has been linked to many health related benefits including maintenance of lean body mass as a result of increased lipolysis (2,11,30) and skeletal muscle adaptations (3,12). It is well known that exercise is a potent stimulator of GH release (6-8,13,18,19,25,28). In young males and females, it has been shown that exercise-induced GH release increases linearly with exercise intensity (13,14,27).

Age, body composition, and fitness level are factors that influence GH release independently and concomitantly (1,4,7,9,10,15,17,21,26,28). The reported significant effect of fitness level on GH release appear to be gender-mediated and equally important as body composition (20,24,25) and, interestingly, several studies have reported the attenuation of GH in overweight and obese individuals regardless of gender (1,5,10). Gender-mediated differences in GH release exist throughout the entire submaximal range of exercise intensities, with more pronounced differences at higher intensities of exercise, although aerobic exercise has been shown to stimulate GH release equally at maximal intensities (28).

Even though both genders display a similar pattern of exercise-induced GH secretion across a broad range of exercise intensities, females tend to release more GH at any given exercise intensity because of greater basal GH secretion (14,25,27,28). During increasing intensities of exercise, females tend to secrete greater GH masses per burst to accommodate for a decrease in the number of GH peaks observed (25,28) during increasing exercise durations beyond 60 min while males typically have a greater GH peak and GH mean value attributed to a greater GH pulse frequency (22). Females have also been shown to obtain a peak GH concentration sooner than males, which is likely due to the fact that they start at a higher basal GH level and reach similar GH maximum exercise values (14,25,27,28). Increasing exercise intensity and duration results in a greater slope in the GH response in males, which is likely to make up for lower basal GH concentrations (26).

We have previously reported no differences in 24-hr IGHC between continuous or intermittent modes of exercise (25) and thus, we used the data from the 30-min continuous constant load exercise session only for this secondary analysis. To our knowledge, no studies have investigated the association of total non-stimulated 24-hr GH concentration with an independent and unrelated acute exercise-induced GH measurement on an unrelated day. Given the known potent role of exercise for stimulating GH release and the numerous beneficial effects of GH on metabolism and body composition, it would be prudent to better understand the relationship between the exercise-induced GH and 24-hr GH concentrations. Information obtained from this study may help clinicians, dieticians and exercise physiologists identify the importance of exercise in the enhancement of GH release and its subsequent metabolic benefits as well as any gender-mediated differences in confounding effects of such variables as gender, age, fitness level, and body composition.

Therefore, the purpose of the current study was to investigate whether the exercise-induced GH response during a single bout of 30-min constant load exercise was related to total 24-hr GH concentration over an unrelated 24-hr period without exercise. In addition, the confounding effects of gender, age, fitness level, and body composition on the relationship between exercise-induced IGHC and total NES 24-hr IGHC during an unrelated day were examined.

## **METHODS**

### **Subjects**

Twenty-nine subjects (N = 16 males and N = 13 females) participated in this study. All subjects were considered sedentary, which was defined as less than two 30-min bouts of exercise per week. Each subject provided written informed consent approved by the University of Virginia

Human Investigators Committee. Each subject was screened for contraindications to exercise and factors known to affect GH secretion, including hematological, renal, hepatic, metabolic, and thyroid function. The subjects were also nonsmokers and not on any systemic medications. All females were pre-menopausal, and none was currently taking contraceptives. The females were tested during the early follicular phase, sometime between day 3 and day 8 after the onset of menses. Descriptive characteristics are presented in Table 1.

## **Procedures**

### ***Maximal Exercise Testing***

Peak oxygen consumption ( $\text{VO}_2$  peak) and blood lactate threshold (LT) were determined from a continuously graded cycle ergometer test (SensorMedics, Yorba Linda, CA) using open circuit spirometry (SensorMedics, Model 2900z, Yorba Linda, CA). Each subject started with an initial workload of 60 watts (W) that was increased by 15 W every 3 min until volitional fatigue occurred. Blood lactate measurements were taken from a forearm vein at rest and during the last 15 sec of each stage (YSI 2700 select biochemistry analyzer; Yellow Springs Instruments, Yellow Springs, OH).  $\text{VO}_2$  peak was chosen as the highest attained  $\text{VO}_2$  value during the test. To determine LT, the blood lactate concentration was plotted against power output and the LT was then chosen as the highest power output obtained before the curvilinear increase in blood lactate concentration occurred.

### ***General Clinical Research Center Admissions***

All subjects refrained from exercise for a 24-hr period prior to being randomly admitted to the General Clinical Research Center (GCRC) on two separate, counterbalanced and randomized occasions. One 24-hr admission included a 30-min bout of constant load exercise and the other 24-hr admission that served as a non-exercise control (NES).

Subjects consumed their evening meal shortly following admission to the GCRC at 5 p.m. Except for the exercise portion of the applicable admission, subjects were asked to remain sedentary. They were only allowed to read, work on computers, and watch television quietly in bed. They were only allowed to be mobile to use the bathroom. At 9 p.m. an intravenous catheter was inserted bilaterally into each forearm vein. At 11 p.m. the lights were turned out. The subjects fasted overnight so that normal nocturnal GH release would not be affected by caloric consumption. The following morning, the subjects either exercised for 30-min on a treadmill (9:00 to 9:30 a.m.) with the exercise intensity set at a constant relative power output midway between the pre-determined  $\text{VO}_2$  at LT and  $\text{VO}_2$  peak (~74%  $\text{VO}_2$  peak, Table 1) or remained sedentary during the other counterbalanced, randomized admission. During this control admission, the subjects were admitted at the same time frame to the GCRC but did not exercise and the total NES 24-hr GH AUC value was used for comparison to the 30-min snapshot of exercise-induced GH release from the separate, unrelated admission.

Subjects were given standardized snack at 8 p.m. the evening before testing (500 kcal, 55% carbohydrate, 15% protein, and 30% fat) as well as standardized meals the following day using strategic intervals at 10 a.m., 2 p.m., and 6 p.m. All meals had identical macronutrient content (55% carbohydrate, 15% protein, and 30% fat). Also, the meals were based on measured basal metabolic rate plus an activity factor for each subject. Each subjects had to consume the entire

meal and the diet, which contained no caffeine or Nutrasweet®. Sodium was controlled at 3 g. All subjects were allowed to consume water ad libitum.

The data from two GCRC admissions was used to examine the relationship between the acute exercise-induced GH response from one randomized admission and 24-hr IGHC on a separate non-related admission where the subject did not exercise. This analysis was done to examine if 24-hr IGHC can be predicted from minimal blood sampling for GH during a 30-min exercise test, along with a host of other common GH predictor variables.

### ***Growth Hormone Measurements***

All measurements were conducted as previously described (25). Briefly, blood samples were taken from an intravenous catheter inserted into a forearm vein every 10 min during exercise and 6 hrs into recovery throughout both admissions. The GH concentrations were measured in duplicate from a single subject by a modified ultrasensitive chemiluminescence assay (Nichols Institute Diagnostics, San Clemente, CA) that had a sensitivity of  $0.005 \mu\text{g}\cdot\text{L}^{-1}$ . Recombinant (22 kDa) human GH was used as the standard (6). Cross-reactivity with 20 kDa recombinant human GH was 30%. Median intra- and interassay coefficients of variation were 5.2% and 6.3%, respectively.

### **Statistical Analyses**

Demographic and subject characteristic data were summarized by the mean (M) and standard deviation (SD) of the measurement distribution. Male versus female comparisons of age, weight, body mass index (BMI), measure of overall physical fitness, and GH responses were examined via a two-sample Wilcoxon rank-sum test.

### ***IGHC***

The 24-hr IGHC was calculated using the trapezoidal integration rule (21). Trapezoidal integration was also used to calculate integrated GH concentration for both 6-hr admissions that included a 30-min bout of exercise. Since exercise alone had such a profound effect on GH release, inclusion of recovery time (1 to 6 hrs) beyond the 30 min of exercise-induced GH during the initial 30-min exercise time frame did not provide additional statistical information about 24-hr IGHC (post-exercise recovery time of 6 hrs only encompassed an additional 5% of total 24-hr IGHC) and thus, only the total 30-min IGHC resulting from the actual exercise session was used in the predictive model of the univariate and multivariate regression analyses along with the exercise-induced peak GH concentration (defined as the highest value attained during the 30 min of exercise).

### ***Univariate and Multiple Regression Analyses***

Within male and female subgroups, Spearman rank correlation analyses were conducted to examine bivariate relationships between the values of the predictor variables. Additionally, the ordinary least-squares (OLS) univariate regression analysis was conducted to examine univariate relationships between 24-hr IGHC and: subject age, BMI, 30-min exercise-induced IGHC, exercise-induced peak GH concentration, and  $\text{VO}_2$  peak. Then, an OLS multivariate regression analysis was conducted to examine covariate adjusted relationships between NES 24-hr IGHC and: subject age, BMI, 30-min exercise-induced IGHC, exercise-induced peak GH concentration, and  $\text{VO}_2$  peak. In order to attain constant residual variance, as required by the OLS regression

model constant variance assumption, the 24-hr IGHC measurements were rescaled to the natural logarithmic scale. Additionally, the natural logarithmic transformation was used to reduce the skewed nature of the 30-min IGHC measurement distribution and to also reduce the skewed nature of the exercise-induced peak GH concentration measurement distribution. With regard to hypothesis testing for the univariate regression analysis, the null hypothesis that there was no predictor variable versus response variable association was tested via the conventional 1 numerator degree freedom F-test for the univariate regression analyses. With regard to hypothesis testing for the multivariate analysis, the null hypothesis that there was no unique association between  $\log_e(24\text{-hr IGHC})$  and the predictor variable was tested via a Type III extra sum of squares F-test, which took into account the variance in the  $\log_e(24\text{-hr IGHC})$  measurements that was explained by the remaining set of predictor variables. A  $P \leq 0.05$  decision rule was used as the null hypothesis rejection criterion. The coefficient of determination  $R^2$  was used to quantify the proportion of the total variation in the  $\log_e(24\text{-hr IGHC})$  measurements that was explained by the values of the predictor variable for both regression models. Spotfire Splus version 8.2 (TIBCO, Inc, Palo Alto, CA) was used to conduct all the statistical analyses.

## RESULTS

Males and females did not differ significantly in age, weight, body mass index (BMI), or  $VO_2$  peak (Table 1). Exercise intensity for the 30-min exercise session was significantly greater in the males versus the females ( $24.7 \pm 1.7$  vs.  $18.7 \pm 2.3$   $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ ,  $P = 0.040$ ). However, the load was relatively adjusted so that each subject exercised at a constant workload equal to  $\sim 74\%$  of  $VO_2$  peak.

**Table 1. Descriptive and Maximal Exercise Testing Data Between Genders.**

Variables	Men	Women	P-value
Age (yrs)	$27 \pm 2$ (18 – 39)	$27 \pm 2$ (17 – 40)	0.979
Weight (kg)	$89.9 \pm 4.3$ (61.1 – 127.7)	$78.4 \pm 5.4$ (54.5 – 123)	0.122
BMI ( $\text{kg} \cdot \text{m}^{-2}$ )	$28.5 \pm 1.7$ (20.2 – 45.8)	$29.1 \pm 2.4$ (20.1 – 46.3)	0.824
$VO_2@LT$ ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ )	$16.0 \pm 1.0$ (9.3 – 22.9)	$12.6 \pm 1.5$ (6.4 – 26.3)	0.060
HR@LT (beats $\cdot \text{min}^{-1}$ )	$122 \pm 6$ (80 – 168)	$126 \pm 3$ (110 – 150)	0.511
Power@LT (W)	$77.2 \pm 4.8$ (35 – 110)	$52.3 \pm 8.8$ (20 – 125)	0.015*
$VO_2$ Peak ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ )	$33.4 \pm 2.5$ (16.4 – 47.4)	$26.3 \pm 3.0$ (13.9 – 46.9)	0.073
HR@ $VO_2$ Peak (beats $\cdot \text{min}^{-1}$ )	$179 \pm 5$ (131 – 213)	$176 \pm 6$ (131 – 200)	0.690
Power@ $VO_2$ Peak (W)	$193.4 \pm 11.8$ (110 – 275)	$131.9 \pm 12.8$ (65 – 215)	0.002
30-Min $VO_2$ ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ )	$24.7 \pm 1.7$ (13.3 – 34.1)	$18.7 \pm 2.3$ (6.9 – 36.6)	0.040*

Means  $\pm$  SD (range); N = 16 males; N = 13 females. LT = lactate threshold. 30-min  $VO_2$  ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ ) was pre-determined exercise intensity quantitatively described as the median between  $VO_2@LT$  ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ ) and  $VO_2$  peak ( $\text{mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ ) in order to keep the relative workloads between genders equal.

\* $P < 0.05$ .

The NES 24-hr IGHC was similar between genders ( $P = 0.488$ ) as well as the 30-min exercise-induced IGHC ( $P = 0.131$ ) and peak GH concentration ( $P = 0.089$ ) (Table 2).

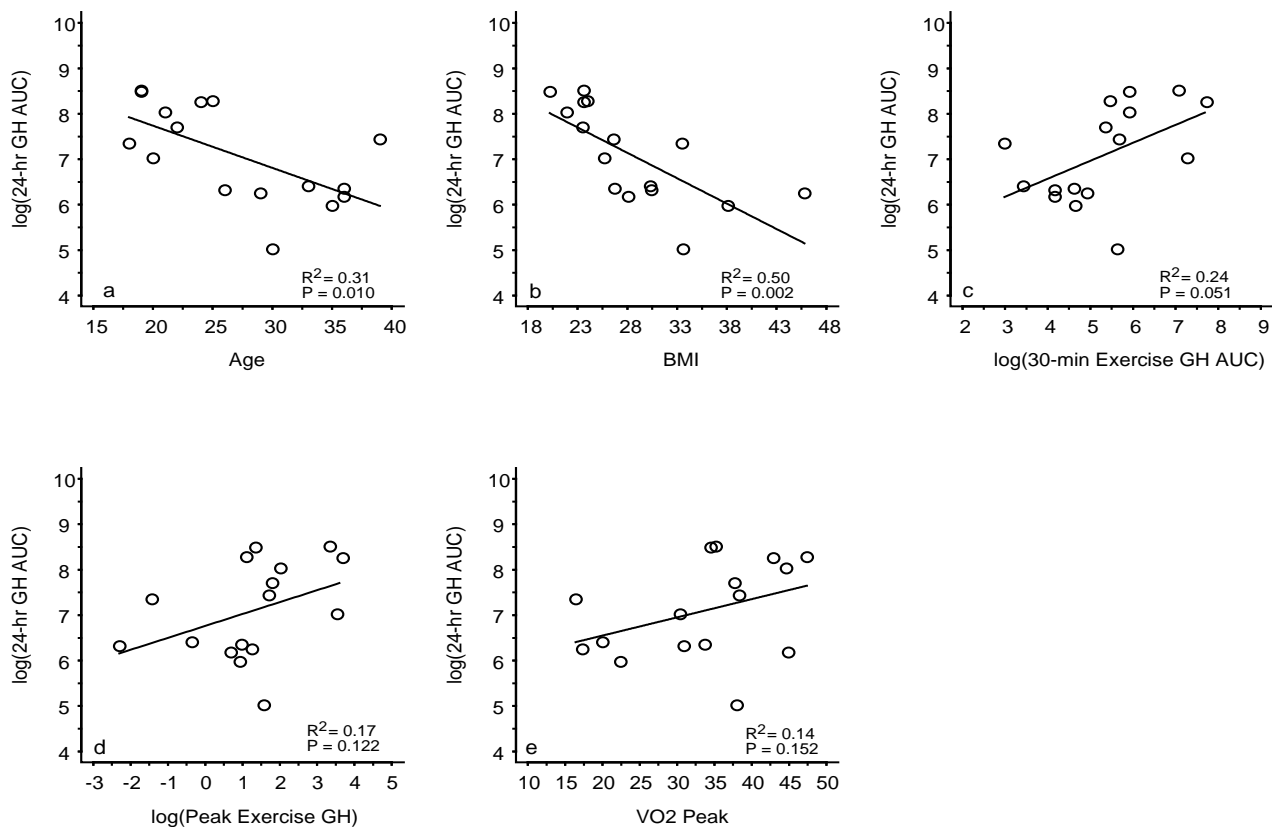
When examining the NES 24-hr IGHC data in males, univariate analyses showed that it was related to BMI ( $r^2 = 0.50$ ,  $P = 0.002$ ) and age ( $r^2 = 0.31$ ,  $P = 0.010$ ), but only marginally with 30-

min of exercise-induced IGHC ( $r^2 = 0.24$ ,  $P = 0.051$ ), and unrelated to exercise-induced peak GH concentration ( $r^2 = 0.17$ ,  $P = 0.122$ ) and  $VO_2$  peak on a separate occasion ( $r^2 = 0.14$ ,  $P = 0.152$ ) (Figure 1).

**Table 2. Total Exercise-Induced GH Concentration and Total Non-Exercise (NES) 24-Hr GH Concentration by Gender.**

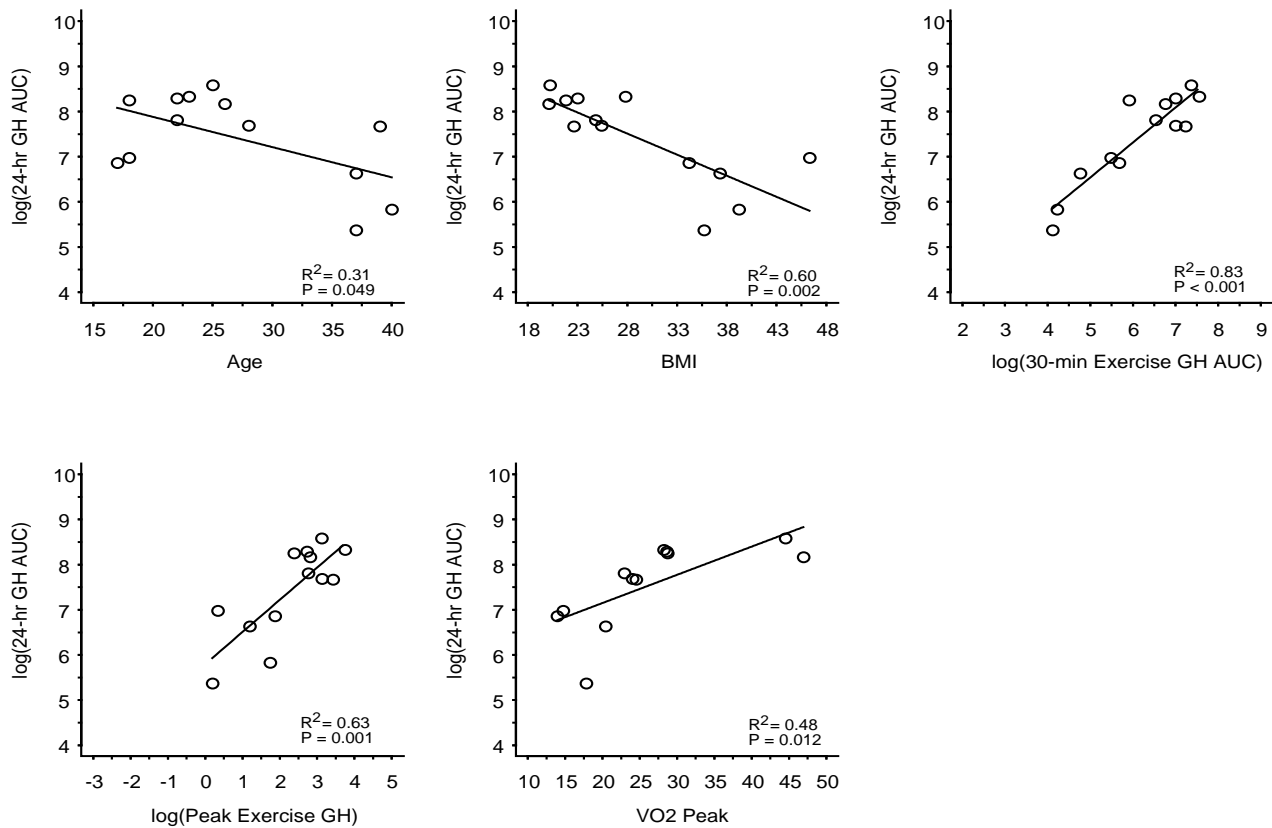
Variables	Men	Women	P-value
<b>30-Min IGHC</b> ( $\mu\text{g}\cdot\text{L}^{-1}\cdot\text{min}^{-1}$ )	448 $\pm$ 158 (20 – 2275)	750 $\pm$ 172 (61 – 1905)	0.131
<b>Peak GH</b> ( $\mu\text{g}\cdot\text{L}^{-1}$ )	9.1 $\pm$ 3.2 (0.1 – 40.4)	15.0 $\pm$ 3.4 (1.2 – 42.4)	0.089
<b>NES 24-Hr IGHC</b> ( $\mu\text{g}\cdot\text{L}^{-1}\cdot\text{min}^{-1}$ )	1891 $\pm$ 418 (150 – 4911)	2353 $\pm$ 455 (212 – 5276)	0.488

Means  $\pm$  SD (range); N = 16 males and N = 13 females.



**Figure 1. Univariate Relationships between  $\log_e(24\text{-Hr GH AUC})$  and: Subject Age, BMI,  $\log_e(30\text{-Min Exercise GH AUC})$ ,  $\log_e(\text{Peak Exercise GH})$  and  $VO_2$  Peak, Within the Male Subgroup.**  $R^2$  denotes the proportion of the total variation in the  $\log_e(24\text{-hr GH AUC})$  values that is explained by the predictor variable values, and  $P$  denotes the  $P$ -value for the test of the null hypothesis that there is no association between the  $\log_e(24\text{-hr GH AUC})$  values and the predictor variable values.

In females, significant univariate relationships were found between NES 24-hr IGHC and age ( $r^2 = 0.31$ ,  $P = 0.049$ ), BMI ( $r^2 = 0.60$ ,  $P = 0.002$ ), 30-min of exercise-induced IGHC ( $r^2 = 0.83$ ,  $P < 0.001$ ), exercise-induced peak GH concentration ( $r^2 = 0.63$ ,  $P = 0.001$ ), and  $VO_2$  peak on an unrelated admission ( $r^2 = 0.48$ ,  $P = 0.012$ ) (Figure 2).



**Figure 2. Univariate Relationships between  $\log_e(24\text{-Hr GH AUC})$  and: Subject Age, BMI,  $\log_e(30\text{-Min Exercise GH AUC})$ ,  $\log_e(\text{Peak Exercise GH})$  and  $VO_2$  Peak, Within the Female Subgroup.**  $R^2$  denotes the proportion of the total variation in the  $\log_e(24\text{-hr GH AUC})$  values that is explained by the predictor variable values, and  $P$  denotes the  $P$ -value for the test of the null hypothesis that there is no association between the  $\log_e(24\text{-hr GH AUC})$  values and the predictor variable values.

In males, the multivariate regression analysis revealed that significant variability in NES  $\log_e(24\text{-hr IGHC})$  was explained by subject age and BMI together ( $P = 0.020$ ), but not by either of the two variables alone ( $P = 0.147$  and  $P = 0.078$ , respectively) after accounting for the variability in NES 24-hr IGHC that was explained by the remaining set of predictor variables. Neither the GH response to exercise (30-min of exercise-induced IGHC, peak exercise-induced GH concentration) or the fitness level (i.e.,  $VO_2$  peak) explained significant unique variability in NES  $\log_e(24\text{-hr GH IGHC})$  ( $P = 0.958$ ,  $P = 0.699$ , and  $P = 0.725$ , respectively) (refer to Table 3). Individually, none of the regression model predictor variables was uniquely associated with NES  $\log_e(24\text{-hr IGHC})$  (Table 3), which in all likelihood is the consequence of the high degree of correlation between several of the predictor variables in males (Table 4).

**Table 3. Multivariate Regression Model for Examining Partial Associations between NES 24-Hr IGHC and Age, BMI, Exercise-Induced IGHC during 30 Min of Exercise, Peak Exercise-Induced GH Concentration and VO<sub>2</sub> Peak in Males.**

Source Of Variation	Parameter Estimate	Standard Error	95% Confidence Interval	F Statistic*	P-value	P-value†
Intercept	11.65	2.71	[5.61, 17.68]			
Age	-0.06	0.04	[-0.14, 0.02]	2.47	0.147	} 0.020
BMI	-0.09	0.05	[-0.20, 0.01]	3.84	0.078	
log <sub>e</sub> (30-Min IGHC)‡	-0.02	0.41	[-0.93, 0.89]	0.00	0.958	} 0.750
log <sub>e</sub> (Peak GH)§	0.12	0.30	[-0.54, 0.78]	0.16	0.699	
VO <sub>2</sub> Peak	-0.01	0.03	[-0.09, 0.06]	0.13	0.725	
Model				4.00	0.030	
R <sup>2</sup> adjusted	0.50					

\*Type III F-statistic for testing the null hypothesis that the “source of variation” is not uniquely associated with log<sub>e</sub>(NES 24-hr IGHC) after taking into account the variation explained by the remaining sources of variation. †P-value of the tests of the null hypothesis that the regression coefficients associated with the two sources of variation are both equal to zero. ‡The natural log transformed values of exercise-induced IGHC during 30 min of exercise. §The natural log transformed values of peak exercise-induced GH concentration. || Adjusted coefficient of multiple determination.

**Table 4. Spearman Correlations between the Individual Predictor Variables of NES 24-Hr IGHC in Males.**

Predictor Variable	Predictor Variable			
	BMI	log <sub>e</sub> (30-Min IGHC)*	log <sub>e</sub> (Peak GH)†	VO <sub>2</sub> Peak
Age	0.47	-0.35	-0.33	0.15
<i>p</i> ‡	0.070	0.180	0.220	0.580
BMI		-0.63	-0.61	-0.55
<i>p</i>		0.010	0.010	0.030
log <sub>e</sub> (30-Min IGHC)*			0.93	0.44
<i>p</i>			<0.001	0.090
Peak GH				0.40
<i>p</i>				0.120

\*Natural log transformed values of exercise-induced IGHC during 30 min of exercise. †Natural log transformed values of peak exercise-induced GH concentration. ‡Denotes the P-value for the test of the null hypothesis that the correlation between the ordered ranks of the two predictor variables is equal to zero.

In the females, multivariate regression analyses revealed that significant unique variability in NES log<sub>e</sub>(24-hr GH IGHC) was explained by the subjects’ age (P = 0.048). The variability in NES log<sub>e</sub>(24-hr IGHC) explained by the GH response to exercise (30 min of exercise-induced IGHC and peak exercise-induced GH concentration) was also significant (P = 0.046), although neither 30 min of exercise-induced IGHC (P = 0.376) or peak exercise-induced GH concentration (P = 0.633) explained unique variability in NES log<sub>e</sub>(24-hr GH IGHC) by itself after accounting for the variability in NES log<sub>e</sub>(24-hr GH IGHC) that was explained by the



remaining set of predictor variables. Like in males, fitness level (i.e., VO<sub>2</sub> peak) was not uniquely associated with NES log<sub>e</sub>(24-hr GH IGHC). Individually, “subject age” was the only predictor variable that was uniquely associated with NES log<sub>e</sub>(24-hr GH IGHC) (Table 5), which again in all likelihood is the consequence of the high degree of correlation between several of the predictor variables (Table 6).

**Table 5. Multivariate Regression Model for Examining Partial Associations between NES 24-Hr IGHC and Age, BMI, Exercise-Induced IGHC during 30 Min of Exercise, Peak Exercise-Induced GH Concentration and VO<sub>2</sub> Peak in Females.**

Source Of Variation	Parameter Estimate	Standard Error	95% Confidence Interval	F Statistic*	P-value	P-value†
Intercept	5.49	1.73	[1.25, 9.73]			
Age	-0.04	0.02	[-0.08, 0.00]	6.17	0.048	} 0.167
BMI	-0.01	0.03	[-0.09, 0.08]	0.02	0.881	
log <sub>e</sub> (30-Min IGHC)‡	0.33	0.34	[-0.51, 1.17]	0.91	0.376	} 0.046
log <sub>e</sub> (Peak GH)§	0.20	0.39	[-0.76, 1.15]	0.25	0.633	
VO <sub>2</sub> Peak	0.03	0.02	[0.02, 0.07]	1.95	0.212	
Model				13.48	0.003	
R <sup>2</sup> adjusted	0.85					

\*Type III F-statistic for testing the null hypothesis that the “source of variation” is not uniquely associated with log<sub>e</sub>(NES 24-hr IGHC) after taking into account the variation explained by the remaining sources of variation. †P-value of the tests of the null hypothesis that the regression coefficients associated with the two sources of variation are both equal to zero. ‡The natural log transformed values of exercise-induced IGHC during 30 min of exercise. §The natural log transformed values of peak exercise-induced GH concentration. || Adjusted coefficient of multiple determination.

**Table 6. Spearman Correlations between the Individual Predictor Variables of NES 24-Hr IGHC in Females.**

Predictor Variable	Predictor Variable			
	BMI	log <sub>e</sub> (30-Min IGHC)*	log <sub>e</sub> (Peak GH)†	VO <sub>2</sub> Peak
Age	0.10	-0.13	0.07	0.14
<i>p</i> ‡	0.730	0.670	0.820	0.670
BMI		-0.66	-0.64	-0.89
<i>p</i>		0.010	0.020	<0.001
log <sub>e</sub> (30-Min IGHC)*			0.92	0.66
<i>p</i>			<0.001	0.020
Peak GH				0.58
<i>p</i>				0.050

\*Natural log transformed values of exercise-induced IGHC during 30 min of exercise. †Natural log transformed values of peak exercise-induced GH concentration. ‡Denotes the P-value for the test of the null hypothesis that the correlation between the ordered ranks of the two predictor variables is equal to zero.

## **DISCUSSION**

The findings of the present study highlight the gender-mediate differences in the relationship between exercise-induced GH release and spontaneous, non-stimulated GH release on a separate day. Our data demonstrate that the IGHC from a 30-min bout of exercise can predict total non-stimulated 24-hr IGHC on a separate, unrelated day in females only. This information should help promote new insight into program design (i.e., diet, exercise, and healthy living habits) for females by clinicians, dieticians, and exercise physiologists. Clearly, the data indicate that exercise is more highly correlated to GH output in females versus males and, therefore, exercise interventions should be the primary focus for achieving optimal health and well-being in females.

While the results are similar to previously reported observations that the GH released during rest and exercise is independently influenced by gender, age, BMI, and physical fitness status, the present study appears to be the first to report gender differences in the joint effects of descriptive predictor variables (age, gender, BMI, and fitness level) in conjunction with the exercise-induced GH response on total 24-hr non-exercise IGHC. Overall, our multiple regression models revealed that 92% of the variance in predicting 24-hr GH IGHC could be explained from age, BMI, fitness level, and 30 min of exercise-induced IGHC in the females compared to 71% in the males. Although the overall regression models explained a significant amount of total variance in both genders, our results provide new evidence of the unique predictor variable differences between males and females.

The primary purpose of the present study was to test the hypothesis that the exercise-induced GH response during an acute 30-min constant load exercise is associated with total non-stimulated 24-hr GH concentration on a separate independent day, regardless of gender, age, BMI, or  $\text{VO}_2$  peak. However, the present findings in the present study indicate that the exercise-induced GH response is associated with and may be used to predict non-stimulated 24-hr IGHC in females but not in males. This finding confirms the gender disparity in the exercise-induced GH response that has been noted in many studies using a variety of exercise modalities, protocols, and measurement techniques. Thus, the finding has practical implications for clinical research and testing especially in regards to exercise interventions aimed at improving total GH output in females.

### **Equating Workload between Genders**

This study used an exercise stimulus that was relative for each subject (power output set midway between the pre-determined  $\text{VO}_2$  at LT and  $\text{VO}_2$  peak), yet similar in intensity to previous research. Despite equating workload intensity between males and females, females had ~1.7 times the amount of exercise-induced IGHC and ~1.6 fold higher exercise-induced peak GH than males during the 30-min constant load exercise bout. There were no statistically significant gender differences in total exercise-induced GH IGHC during the 30-min constant load exercise ( $P = 0.131$ ), the peak GH during exercise ( $P = 0.089$ ), or the unrelated non-exercise 24-hr IGHC ( $P = 0.488$ ), but results from the multivariate regression analysis showed that the exercise-induced GH response still represented a greater portion of the total 24-hr GH output on a separate, unrelated non-exercise 24-hr occasion in females versus males. Total GH concentration during 30 min of constant load exercise accounted for ~31.8% of the total 24-hr GH

concentration in females compared to ~23.7% in males. Therefore, total non-exercise 24-hr GH concentration can be adequately predicted by the exercise-induced GH response in as little as 30 min of exercise on a separate, unrelated occasion in females only (Table 5).

### **Exercise-Induced Growth Hormone “Snapshot”**

Even though 30 min is 2% of a 24-hr period, a substantial amount of GH was released during that short time frame (specifically, 23.7% in males and 31.8% in females). Therefore, when adding additional recovery time (330 min) that would explain 27% of a 24-hr period, the total GH concentration released increased minimally to 28.4% and 36.1% in males and females, respectively of the total 24-hr GH concentration. Based on negligible 5% differences in GH concentration and no associated changes in the statistical model when more post-exercise recovery time was included in the analysis, we conclude that minimal additional information related specifically to the total 24-hr GH concentration is gleaned from collecting GH profiles during recovery from exercise. However, monitoring GH concentration during the post-exercise recovery period may also provide valuable information about other incremental GH secretion variables (e.g., time to return to basal secretion, GH half-life, etc.).

### **Fitness Level and Growth Hormone Release**

Additional parameters known to concomitantly influence total GH concentration were examined in an attempt to further explain the association between exercise-induced and total non-exercise 24-hr GH concentration. Peak oxygen consumption ( $VO_2$  peak) was significantly related to non-exercise 24-hr IGHC in females (Figure 2). But, when using the multivariate regression analysis to examine the covariate adjusted relationship,  $VO_2$  peak was no longer uniquely associated with total non-exercise 24-hr IGHC for either gender. Weltman et al. (24) previously found that fitness level (as assessed by  $VO_2$  peak) was significantly correlated with total 24-hr GH concentration in males but not females.

Although the subjects used in this study are slightly more overweight and less fit, further examination of the regression line from Figure 1d of Weltman et al. (24) suggests that at lower  $VO_2$  peak values (such as  $<40 \text{ mL} \cdot \text{kg}^{-1} \cdot \text{min}^{-1}$ ) the relationship between  $VO_2$  peak and total 24-hr GH concentration is stronger in females than males. If this interpretation is accurate, then further work needs to be done across broader ranges of exercise capacity to determine if  $VO_2$  peak is truly a significant predictor of 24-hr GH release. This also suggests that  $VO_2$  peak is not the only determinant of exercise capacity and our data on the “snapshot” of 30 min of constant load total exercise-induced GH concentration may provide unique input on total 24-hr GH concentration that is especially true in females.

### **BMI, Gender, and Growth Hormone Release**

In a previous study, total 24-hr GH concentration was simultaneously determined by BMI and gender, while peak GH values were significantly influenced by BMI in both males and females (23). In the current study, multiple regression analysis revealed that significant variability in non-exercise 24-hr IGHC was explained by age and BMI in males but not independently after accounting for the variability that was explained by the other predictor variables (Table 3). However, in females, only age was uniquely associated with non-exercise 24-hr IGHC (Table 5). Although the range of BMI in the current study was large and could have influenced the results

of the statistical model, it was similar to the BMI range reported by Veldhuis et al. (23). While they also reported BMI to have a significant negative relationship to total GH concentration across both genders, it was only significant when jointly associated with gender and IGF-1 simultaneously. An examination of the results of Weltman et al. (24) suggest that: (a) at lower BMI levels the relationship between BMI and GH concentration appears to be more strongly associated in the males compared to the females; and (b) at higher BMI levels the relationship between BMI and GH concentration may become stronger in the females.

In obese individuals, many of the beneficial effects of GH are absent, regardless of gender or age in this population (9,22). The lack of a relationship between BMI and GH concentration in females using the multivariate statistical approach is unexpected, considering GH has a strong lipolytic effect (11,13) and women typically release greater amounts of GH. Clasey et al. (1) conducted a large-scale study on physiological factors known to alter 24-hr GH secretion. They concluded that abdominal visceral fat was a stronger predictor of 24-hr pulsatile GH secretion than BMI, regardless of age or gender. An individual's BMI does not distinguish between lean and fat mass and phenotypic differences in body composition known to be gender-mediated. In general, females accumulate more subcutaneous fat than visceral fat compared to males and males have more lean muscle mass in their arms while females tend to accumulate more of their fat mass in the lower body (29). However, it is not known whether total exercise-induced GH concentration may be influenced by gender differences in subcutaneous versus intra-visceral fat reserves during exercise. Regardless, the biological activity of GH indicates that exercise may be an important precursor to beneficial changes in lean body mass and body fat. The GH has been shown to augment adipose tissue lipolysis and, therefore, exercise-induced GH release may be a good therapeutic agent for obtaining optimal body composition (especially in females).

### **Practical Laboratory Testing Implications**

Laboratory testing procedures to determine total 24-hr IGHC are time consuming and costly and also require the use of a highly sensitive assay (16). Stimulation tests can report the level of GH in the blood after the infusion of GH stimulants (i.e., insulin, glucagon, arginine or GH-releasing hormone) that measure the ability of the pituitary gland to release GH. Even though these tests can be done in shorter time frames than 24 hrs, these tests can still take several hours. Given the significant participant and laboratory burden related to the collection of time-lengthened GH profiles to examine 24-hr GH output, exercise-induced GH results could provide researchers with the capability of using a brief exercise-induced snapshot of the GH response to estimate total 24-hr GH concentration in females without having to conduct extensive 24-hr clinical testing.

### **Limitations**

Our statistical regression model used 16 males with 2 predictors, which gave an 8:1 subject-to-predictor. Our 14 females to 3 predictor ratio is less than the suggested 10:1 subject-to-predictor ratio for running continuous variables. Therefore, we recognize the slight limitation of statistical power with our methodology. Additionally, the low fitness level of our subjects (10th percentile) means we can only infer our results on the "un-fit" population and more fit subjects may not respond in exactly the same manner. Since age has a unique association with GH release, our results also cannot be interpreted for older adult populations. Regardless of the possibility of over-fitting our data and our focused demographics, the exact mechanisms causing gender

differences in the relationship between the exercise-induced GH response during an acute 30-min bout of exercise and total 24-hr GH concentration are unknown and our results provide valuable information for the young adult to middle-aged “un-fit” population.

## CONCLUSIONS

To our knowledge, this is the first study to use constant load exercise-induced GH concentration in a multivariate regression analysis to predict total non-exercise 24-hr GH output on a separate occasion in human subjects. While exercise-induced GH concentration during constant load exercise (~74 VO<sub>2</sub> peak) provided important information on total non-exercise 24-hr IGHC in females on a separate occasion, it had no relationship to total non-exercise 24-hr GH concentration in males. Thus, we conclude that total exercise-induced GH concentration is a stronger predictor of total non-stimulated 24-hr GH concentration in females and those interventions for increasing total 24-hr GH concentration should be aimed at increasing exercise habits in females.

Additionally, GH release during a 30-min constant load exercise can provide valuable insight into total non-exercise 24-hr GH output in females on a separate occasion without extensive 24-hr blood sampling. Gender disparities in GH release are well documented at rest, but less is known about the gender differences in exercise-induced GH concentration and how these differences may influence total 24-hr GH concentration. While many GH stimulation tests are available for the diagnosis of GH deficiencies, the exercise-induced GH response to constant load exercise may provide additional insights into clinically relevant subtle gender differences in GH release.

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## REFERENCES

1. Clasey JL, Weltman A, Patrie J, Weltman JY, Pezzola S, Bouchard C, Thorner MO, Hartman ML. Abdominal visceral fat and fasting insulin are important predictors of 24-hour GH release independent of age, gender, and other physiological factors. *J Clin Endocrinol Metab.* 2001; 86:3845-3852.
2. Djurhuss CB, Gravholt CH, Nielsen S, Pedersen SB, Moller N, Schmitz O. Additive effects of cortisol and growth hormone on regional and systemic lipolysis in humans. *Am J Physiol.* 2004;286:EE488-EE494.
3. Doessing S, Heinemeier KM, Holm L, Mackey AL, Schjerling P, Rennie M, Smith K, Reitelseder S, Kappelgaard AM, Rasmussen NH, Flyvbjerg A, Kjaer M. Growth

hormone stimulates the collagen synthesis in human tendon and skeletal muscle without affecting myofibrillar protein synthesis. *J Physiol.* 2010;588:341-351.

4. Ho KY, Evans WS, Blizzard RM, Veldhuis JD, Merriam GR, Samojilk E, Furlanetto R, Rogol AD, Kaiser DL, Thorner MO. Effects of sex and age on the 24-hour profile of growth hormone secretion in man: Importance of endogenous estradiol concentrations. *J Clin Endocrinol Metab.* 1987;64:51-58.
5. Irving BA, Weltman JY, Patrie JT, Davis Ck, Brock DW, Swift D, Barrett EJ, Gaesser GA, Weltman A. Effects of exercise training intensity on nocturnal growth hormone secretion in obese adults with the metabolic syndrome. *J Clin Endocrinol Metab.* 2009;94:1979-1986.
6. Kanaley JA, Weatherup-Dentes MM, Jaynes EB, Hartman ML. Obesity attenuates the growth hormone response to exercise. *J Clin Endocrinol Meta.* 1999;84:3156-3161.
7. Kanaley JA, Weltman JY, Veldhuis JD, Rogol AD, Hartman ML, Weltman A. Human growth hormone response to repeated bouts of aerobic exercise. *J Appl Physiol.* 1997;1756-1761.
8. Kindermann W, Schnabel A, Schmitt WM, Biro G, Cassens J, Weber F. Catecholamines, growth hormone, cortisol, insulin, and sex hormones in anaerobic and aerobic exercise. *Eur J Appl Physiol.* 1982;49:389-399.
9. Lin CM, Huang YL, Lin ZY. Influence of gender on serum growth hormone, insulin-like growth factor-I and its binding protein-3 during aging. *Yonsei Med J.* 2009;50:407-413.
10. Makimura H, Stanley T, Mun D, You SM, Grinspoon S. The effects of central adiposity on growth hormone (GH) response to GH-releasing hormone-arginine stimulation testing in men. *J Clin Endocrinol Metab.* 2008 93:4254-4260.
11. Nass R, Thorner MO. Impact of GH-cortisol ratio on the age-dependent changes in body composition. *Growth Horm IGF Res.* 2002;12:147-161
12. Nindl BC, Scoville CR, Sheehan KM, Leone CD, Mello RP. Gender differences in regional body composition and somatotrophic influences of IGF-1 and leptin. *J Appl Physiol.* 2002;92: 1611-1618.
13. Prifzlaff CJ, Wideman L, Blumer J, Jensen M, Abbott RD, Gaesser GA, Veldhuis JD, Weltman A. Catecholamine release, growth hormone secretion, and energy expenditure during exercise vs. recovery in men. *J Appl Physiol.* 2000;89:937-946.
14. Pritzlaff CJ, Wideman L, Weltman JY, Abbott R, Gutgesell M, Hartman ML, Veldhuis JD, Weltman A. Gender governs the relationship between exercise intensity and growth hormone release in young adults. *J Appl Physiol.* 2002;92:2053-2060.

15. Quabbe HJ, Bratzke HJ, Siegers U, Elban K. Studies on the relationship between plasma free fatty acids and growth hormone secretion in man. *J Clin Invest.* 1972;51:2388-2398.
16. Reutens AT, Hoffman DM, Leung KC, Ho KY. Evaluation and application of a highly sensitive assay for serum growth hormone (GH) in the study of adult GH deficiency. *J Clin Endocrinol Metab.* 1995;80(2):480-485.
17. Salvadori A, Fanari P, Marzullo P, Codecasa F, Tovaglieri L, Cornacchia M, Walker G, Brunani A, Longhini E. Dynamics of GH secretion during incremental exercise in obesity, before and after a short period of training at different work-loads. *Clin Endocrinol.* 2010;73:491-496.
18. Sartorio A, Agosti F, Marinone PG, Proietti M, Lafortuna CL, Maffiuletti NA. Growth hormone responses to repeated bouts of aerobic exercise with different recovery intervals. *J Appl Physiol.* 2006;100:1093-1094.
19. Stokes KA, Nevill ME, Lakomy HK, Hall GM. Reproducibility of the growth hormone response to sprint exercise. *Growth Horm IGF Res.* 2003;13:336-340.
20. Vahl N, Jergensen JO, Skjaerbaek C, Veldhuis JD, Orskov H, Christiansen JS. Abdominal adiposity rather than age and sex predicts mass and regularity of GH secretion in healthy adults. *Am J Physiol.* 1997;272(6):E1108-1116.
21. Veldhuis JD, Johnson ML. Cluster analysis: A simple, versatile, and robust algorithm for endocrine pulse detection. *Am J Physiol.* 1986;250:E486-E493.
22. Veldhuis JD, Patrie JT, Brill KT, Weltman JY, Mueller EE, Bowers CY, Weltman A. Contributions of gender and systemic estradiol and testosterone concentrations to maximal secretagogue drive of burst-like growth hormone secretion in healthy middle-aged and older adults. *J Clin Endocrinol Metab.* 2004;89:6291-6296.
23. Veldhuis JD, Roelfsema F, Keenan DM, Pincus S. Gender, age, body mass index, and IGF-1 individually and jointly determine distinct GH dynamics: Analyses in one hundred healthy adults. *J Clin Endocrinol Metab.* 2011;96:115-121.
24. Weltman A, Weltman YJ, Hartmann ML, Abbott RD, Rogol AD, Evans WS, Veldhuis JD. Relationship between age, percentage body fat, fitness, and 24-hour growth hormone release in healthy young adults: Effects of gender. *J Clin Endocrinol Metab.* 1994;78(3):543-548.
25. Weltman A, Weltman YJ, Waston-Winfield DD, Frick K, Patrie J, Kok P, Keenan DM, Gaesser GA, Veldhuis JD. Effects of continuous versus intermittent exercise, obesity, and gender on growth hormone secretion. *J Clin Endocrinol Metab.* 2008;93:4711-4720.

26. Wideman L, Consitt L, Patrie J, Swearingin B, Bloomer R, Davis P, Weltman A. The impact of sex and exercise duration on growth hormone secretion. *J Appl Physiol.* 2006;101:1641-1647.
27. Wideman L, Weltman Y, Patrie J, Bowers CY, Shah N, Story S, Weltman A, Veldhuis JD. Synergy of L-arginine and growth hormone (GH)-releasing peptide-2 on GH release: Influence of gender. *Am J Physiol.* 2000;279:R1455-1466.
28. Wideman L, Weltman JY, Shah N, Story S, Veldhuis JD, Weltman A. Effects of gender on exercise-induced growth hormone release. *J Appl Physiol.* 1999;87:1154-1162.
29. Wirth A, Steinmetz B. Gender differences in changes in subcutaneous and intra-abdominal fat during weight reduction: An ultrasound study. *Obes Res.* 1998;6:393-399.
30. Zhao JY, Cowley MJ, Lee P, Birzniece V, Kaplan W, Ho KK. Identification of novel GH-regulated pathway lipid metabolism in adipose tissue: A gene expression study in hypopituitary men. *J Clin Endocrinol Metab.* 2011;96:1188-1196.

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