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**Establishment of optimal control strategies
to eliminate bovine viral diarrhoea
in New Zealand**

A thesis presented in partial fulfilment of the requirements for the degree of

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Abstract

Although there has been a noticeable reduction in the prevalence of bovine viral diarrhoea virus (BVDV) in New Zealand over the past decade, it is well recognised that New Zealand will need a systematic compulsory programme to eliminate BVDV. The aims of this thesis were to address the knowledge gaps around the epidemiology and economics of BVDV to explore the cost-effectiveness of different national BVDV control frameworks.

First, the risk factors for BVDV infection were explored using data collected from cattle herds across the country. A Bayesian network analysis revealed that animal contacts between neighbouring farms significantly increased the risk of herds being seropositive for BVDV. The second study used data collected from New Zealand commercial beef farms to estimate the transmission rate of BVDV from extensively grazed persistently infected (PI) animals. Using an approximate Bayesian computation method, the BVDV transmission rate was estimated at 0.11 per PI animal per day, which was lower than previously derived estimates for dairy herds and intensively farmed beef herds.

For the third study, BVDV simulation models were developed for New Zealand dairy and beef farms to estimate the economic impacts of BVDV outbreak and to identify the most cost-effective control strategies at an individual farm level. The direct losses due to BVDV outbreak were estimated as NZ\$ 22.22 per dairy cow per year and NZ\$ 41.19 per beef cow per year. Annual testing to cull identified PI calves and annual vaccination were economically beneficial to control a BVDV outbreak for a dairy and beef breeding farm, respectively. In the fourth study, BVDV transmission was simulated at a national scale with the models, predicting that BVDV could be successfully and economically controlled by requiring dairy farms to double fence boundaries and perform either annual calf testing or herd-level screening test and requiring beef farms to conduct annual vaccination.

Overall, the findings from the thesis highlight that BVDV elimination is both technically feasible and cost-efficient in New Zealand. The outputs of this thesis can be used to

facilitate discussion with farmers and stakeholders about the benefits and feasibility of national BVDV elimination in New Zealand.

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Table of Contents

Abstract.....	iii
Acknowledgements.....	v
List of Publications.....	vii
List of Presentations.....	viii
Table of Contents.....	ix
List of Figures.....	xiii
List of Tables.....	xvi
Chapter 1.....	1
1.1. Epidemiological features of BVDV.....	1
1.2. Global situation.....	1
1.3. BVDV in New Zealand.....	3
1.4. BVDV elimination in New Zealand.....	3
1.5. Thesis structure.....	5
Chapter 2.....	7
2.1. Abstract.....	7
2.2. Introduction.....	8
2.3. General features of BVD.....	9
2.3.1. Biology of BVD.....	9
2.3.2. Transmission.....	11
2.3.3. Diagnosis.....	12
2.3.4. Vaccination.....	14
2.4. BVD in New Zealand.....	15
2.4.1. Prevalence of BVD.....	15
2.4.1.1. Prevalence in the dairy sector.....	15
2.4.1.2. Prevalence in the beef sector.....	16
2.4.1.3. Within-herd prevalence of PI animals.....	17
2.4.2. Risk factors for transmission.....	17
2.4.3. Financial losses due to BVD.....	18
2.5. Control of BVD.....	19
2.5.1. European programmes.....	19
2.5.1.1. Scandinavian framework.....	19
2.5.1.2. Swiss framework.....	20
2.5.1.3. German framework.....	20

2.5.2. Validity of diagnostic tests in New Zealand	21
2.5.3. Validity of BVD vaccination in New Zealand	22
2.5.4. Financial benefit of BVD control programmes	23
2.6. Discussion	24
2.7. Acknowledgements	27
Chapter 3	29
3.1. Abstract	29
3.2. Introduction	30
3.3. Material and methods	32
3.3.1. Data collection	32
3.3.2. Data processing	34
3.3.3. Descriptive statistics	35
3.3.4. Multivariable logistic regression	36
3.3.5. BN modelling	37
3.4. Results	38
3.4.1. Descriptive statistics from the survey responses	38
3.4.2. Multivariable regression analysis	45
3.4.3. Bayesian network analysis	47
3.5. Discussion	48
3.6. Conclusions	53
3.7. Acknowledgement	53
Chapter 4	55
4.1. Abstract	55
4.2. Introduction	56
4.3. Materials and methods	59
4.3.1. Data collection and extraction	59
4.3.2. Simulation model	60
4.3.3. Estimation of within-herd BVDV transmission rates	66
4.3.4. Validation of the estimated parameters	68
4.4. Results	69
4.4.1. Longitudinal data	69
4.4.2. Parameters estimation	70
4.4.3. Validation	70
4.5. Discussion	72
4.6. Conclusions	77
4.7. Acknowledgements	77

Chapter 5.....	79
5.1. Abstract.....	79
5.2. Introduction.....	80
5.3. Material and methods.....	82
5.3.1. Demographics of dairy farms.....	82
5.3.2. Demographics of beef farms.....	88
5.3.3. Disease component.....	89
5.3.4. Production and economic component.....	96
5.3.5. BVDV control strategies.....	97
5.3.6. Model initiation and BVDV outbreak.....	99
5.3.7. BVDV dynamics and economic analysis.....	100
5.3.8. Sensitivity analysis.....	102
5.4. Results.....	105
5.4.1. Dynamics of BVDV and its economic impact.....	105
5.4.2. Cost-benefit of BVDV control strategies.....	105
5.4.3. Sensitivity analysis.....	110
5.5. Discussion.....	110
5.6. Conclusion.....	116
Acknowledgements.....	117
Chapter 6.....	119
6.1. Abstract.....	119
6.2. Introduction.....	120
6.3. Material and methods.....	123
6.3.1. Data processing.....	123
6.3.1.1. Number of cattle farms.....	126
6.3.1.2. Herd size.....	127
6.3.1.3. Planned start of mating.....	128
6.3.1.4. Adjacent farms.....	129
6.3.1.5. Cattle movements.....	130
6.3.2. National model.....	132
6.3.2.1. Demographics of cattle within a farm.....	132
6.3.2.2. Movement reconstruction.....	137
6.3.2.3. BVDV transmission within a farm.....	138
6.3.2.4. BVDV transmission between farms.....	139
6.3.2.5. Model initiation.....	139
6.3.3. National BVDV control programmes.....	140

6.3.3.1. Movement restriction	141
6.3.3.2. Annual testing and culling.....	141
6.3.3.3. Annual screening test	141
6.3.3.4. Vaccination.....	142
6.3.3.5. Double-fencing.....	142
6.3.4. Model Outputs.....	143
6.4. Results.....	144
6.4.1. Effectiveness of national BVDV control programmes.....	144
6.4.2. Cost-benefit analysis of national BVDV control programmes.....	150
6.5. Discussion	151
6.6. Conclusions	156
6.7. Acknowledgements	156
Chapter 7	158
7.1. Overview of key findings	158
7.2. Risk factors for BVDV infection in New Zealand cattle farms	159
7.3. BVDV dynamics in extensively grazed herds.....	162
7.4. Economics of BVDV and its control.....	163
7.5. National BVDV control programmes	165
7.6. Methodological limitations	167
7.7. Future work: feasibility of BVDV elimination in New Zealand.....	170
7.7.1. Accurate data.....	170
7.7.2. Social acceptability	171
7.7.3. Governmental support	172
7.7.4. Other possible limitations.....	173
7.8. Conclusion.....	174
Bibliography.....	176
Appendix 1	203
Appendix 2	206
Appendix 3	236
Appendix 4	246

List of Figures

Figure 3. 1. Figure 3.1. Proportion of cattle farms actively infected with BVD. Note that no cattle farms from three regions (Gisborne, Taranaki, and Nelson) were sampled. 39

Figure 3. 2. Additive Bayesian network representing significant associations between farm management risk factors on active infection with BVD virus. For each arc, the median odds ratio is reported, and solid and dashed arcs indicate positive and negative associations, respectively. Blue arcs indicate associations with beef farm, and red counterpart describing the one with dairy farms. Active BVD virus infection status is highlighted in grey, and production type (beef and dairy) is described in rectangle node. 46

Figure 4. 1. Description of within-herd simulation for estimating within-herd BVDV transmission rate. Blue and red shaded areas indicate the breeding and calving periods, respectively. Red dotted arrow represents the contribution to the force of BVDV infection. 64

Figure 4. 2. The posterior distribution of within-herd BVDV transmission rates for PI animals (βP). The mode was 0.11 with a 95% highest posterior density interval between 0.03 and 0.34. 71

Figure 4. 3. The posterior distribution of the estimated parameters for each herd. Left: the initial proportion of BVDV seropositive animals (μ), Middle: the proportion of introduced PI animals (ρ), Right: the day of PI animals being introduced (τ). The distributions of the first row illustrate the prior distribution of each parameter. Green and red shaded areas indicate the first breeding and calving period, respectively. Day FS indicates the day of first sampling. 72

Figure 5. 1. Schematic representation of demographic events in a dairy farm. C, YH, BH, and MA indicate calf herd, young replacement heifer herd, breeding replacement heifer herd, and mixed-aged cow herd, respectively. Filled and shaded area in the mating period indicate the period of artificial insemination and natural mating with bulls, respectively. 85

Figure 5. 2. Schematic representation of demographic events in a beef farm. CH, CM, YH, BH, MA, FH, and FS indicate calf herd from breeding heifers, calf herd from mixed-aged cow herd, young heifer herd, breeding heifer herd, mixed-aged cow herd, fattening heifer herd, and fattening steer herd, respectively. 85

Figure 5. 3. Schematic representation of transition between BVDV infection status (top) and BVDV infection status of new born calves (bottom). BVDV infection status consists of passively immunised via maternal antibody (M), susceptible (S), transiently infected (TI), persistently infected (PI), recovered (R), vaccinated insufficiently to prevent fetal infection (VI), and vaccinated sufficiently to prevent fetal infection (VS). Top: Solid black arrows indicate the transition between BVDV infection status, blue dotted arrows indicate the decay of immunity conferred by vaccination, and red dashed arrows illustrate the force of BVDV infection. Bottom: Solid arrows indicate the BVDV infection status of calves depending on the infection status of their dams. Dashed arrows from VI illustrate the BVDV infection status of calves if VI dams were infected. 93

Figure 5. 4. The number of transiently and persistently infected animals during the simulation period in New Zealand dairy (left) and beef (right) farms. 103

Figure 5. 5. The economic impact of BVDV introduction on naïve dairy (left) and beef (right) farms in New Zealand. Green and orange bars indicate the gross farm revenue without and with BVDV, respectively.104

Figure 5. 6. Average annual benefit (green) and cost (red) of BVDV control strategies per mixed-age cow for dairy (left) and beef (right) farms in New Zealand. Non-significant or negative benefits are not presented in this figure, and any average benefit/cost ratio of >1 is written in red. TC, VC, AB, and DF indicate testing and culling, vaccination, annual boosting, and double fencing, respectively. C, S, 1 and 2 indicate the annual testing of female breeding calves, annual screening test, vaccination starting before one year old; and vaccination starting at one year old, respectively.109

Figure 6. 1. Description of data processing. Data between 2014 and 2016 were extracted from the National Animal and Identification Tracing system in New Zealand. Only the data inside blue solid squares were explicitly incorporated into the national model, while the one inside a grey dashed square was indirectly modelled. † indicates the number of batches (i.e. a group of cattle moved from a source farm to a destination farm in a given week). Key: PSM, planned start of mating; U, Uniform distribution (lower, upper limit), N, normal distribution (mean, variance).125

Figure 6. 2. Distribution of cattle farms in New Zealand (left) and the Canterbury region (right). Left & Right: Red and blue dots indicate dairy and beef farms, respectively, with the size of dot was proportional to the mixed-age cow herd size for each farm. Right: Grey lines in the Canterbury region indicate aggregated animal movements via trade during 2014-2016 (direction was ignored for the simplicity of illustration).130

Figure 6. 3. Illustration of demographic events in the New Zealand dairy and beef industries. SD, start of dry-off period; OG, week to move from/to off-site grazing location; PSC, planned start of calving; PSM, planned start of mating; Testing and culling, annual testing and culling of female breeding calves; Screening test, annual screening of bulk tank milk (dairy) or serum samples of 15 young heifers (beef) followed by a PI hunt; VX 1st, 2nd, and 3rd, the first, second, and third BVDV vaccination, respectively; AB, annual boosting of BVDV vaccination.133

Figure 6. 4. National prevalence of BVDV positive farms (black) with respect to implementing no BVDV elimination programme, movement restrictions (P1), and annual testing and culling of breeding female calves (P2) for both dairy and beef breeding farms. Red and blue lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively.146

Figure 6. 5. National prevalence of BVDV positive farms (black) with respect to implementing annual screening test followed by a PI hunt (P3), double fencing (P4), and vaccination of breeding animals with annual boosters (P5) for both dairy and beef breeding farms. Red and blue lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively.147

Figure 6. 6. National prevalence of BVDV positive farms (black) with respect to implementing double fencing (P6), annual testing and culling of breeding female calves with double fencing (P7), and annual screening test followed by a PI hunt with double fencing (P8) for dairy farms. For all strategies, vaccination of breeding animals with annual boosters was implemented for beef farms. Red and blue

lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively. 148

Figure 6. 7. Accumulated benefit (blue) and BVDV control cost (red) for different BVDV elimination programmes over time. Solid line and shaded area indicate median and 95% prediction interval, respectively. 149

List of Tables

Table 3. 1. Description of survey responses and univariable logistic regression of farm management risk factors for active infection with BVD virus from 304 New Zealand cattle herds. Median and interquartile range of number of farms over 1,000 imputed datasets are also presented. Odds ratio with bold and underlined characters shows p-value less than 0.2.....	41
Table 3. 2. Multivariable logistic regression of significant farm management risk factors on active infection with BVD in 304 New Zealand cattle farms.	44
Table 4. 1. Definition and value of parameters used for within-herd BVDV transmission model.	65
Table 4. 2. Prior distribution of parameters estimated using ABC-SMC algorithm.	66
Table 4. 3. Mode (95% highest posterior density interval) of the estimated initial proportion of BVDV seropositive animals (μ), the proportion of introduced PI animals (ρ), and the day of PI animals being introduced (τ) for each studied farm.	71
Table 5. 1. Model parameters for demographic and BVDV components.	86
Table 5. 2. Model parameters for production/economic and control components.	94
Table 5. 3. List of BVDV control options in three different categories. A BVDV control strategy was a combination of three control options from each of category (30 strategies in total).....	102
Table 5. 4. Parameter values used for the sensitivity analysis.	103
Table 5. 5. Benefit and cost of implementing BVDV testing and culling (TC), vaccination (VC) with/without annual boosting (AB), and double fencing (DF) in New Zealand dairy farm. Benefit-cost (B/C) ratios of control strategies with significant benefit are shown in bold.	107
Table 5. 6. Benefit and cost of implementing BVDV testing and culling (TC), vaccination (VC) with/without annual boosting (AB), and double fencing (DF) in New Zealand dairy farm. Benefit-cost (B/C) ratios of control strategies with significant benefit are shown in bold.	108
Table 6. 1. Description of parameter values used in the national BVDV simulation model.....	134
Table 6. 2. Description of the national BVDV control programmes.....	143
Table 6. 3. Median (95% prediction interval) of the % reduction in the number of BVDV positive farms between the beginning and end of simulation period (i.e. % reduction) and the last week that the national prevalence of BVDV positive farms being 2.0% (i.e. last week for 2.0%) with respect to implementing different national BVDV control programmes.....	145
Table 6. 4. Median (95% prediction interval) of accumulated benefit (Benefit), BVDV control cost (Cost), year of break-even point, and benefit/cost ratio with respect to implementing different national BVDV control programmes.	150

Chapter 1

Introduction

1.1. Epidemiological features of BVDV

Bovine viral diarrhoea virus (BVDV) is a *pestivirus* in the *Flaviviridae* family that can have significant detrimental impacts on animal health and production performance (Lanyon et al., 2014). Cattle infected with BVDV have a wide range of subclinical and clinical signs including diarrhoea, respiratory illness, immunosuppression, congenital defects, and abortion/stillbirth (Fray et al., 2000). The key characteristic of BVDV epidemiology is the presence of persistently infected (PI) animals that shed large quantities of the virus throughout their life time (Houe & Meyling, 1991). A PI calf can be generated if a heifer or cow that is naïve to BVDV becomes infected during the risk period of early-to-mid pregnancy and then subsequently gives birth to a live calf. PI calves are generally ill-thrifty and have a short life expectancy, however, apparently normal looking PI calves can survive for more than three years (Hill, Reichel, et al., 2010; Voges et al., 2006). Accordingly, the most important risk factor for naïve cattle to be exposed to BVDV is through direct contacts with PI animals (Houe & Meyling, 1991). In contrast to persistent infections, animals that are transiently infected (TI) with BVDV shed lower levels of the virus for a limited time (Thurmond, 2005) and typically recover within two to three weeks of infection (Duffell & Harkness, 1985).

1.2. Global situation

BVDV is prevalent in most cattle producing countries worldwide (Anonymous, 2020). A meta-analysis of global BVDV prevalence reported an average of 67.7% cattle herds across 73 countries showing evidence of recent BVDV exposure with 27.2% of herds likely containing PI animals (Scharnböck et al., 2018). Although the prevalence of PI animals

within an actively infected farm is generally less than 5% (Richter et al., 2019), previous studies have shown that even small numbers of PI animals on a cattle farm could result in the seroconversion of up to 97% of susceptible animals in the herd in a given year (Houe & Meyling, 1991). Given the high infectious pressure of PI animals and their prevailing presence across cattle industries, BVDV is known to have significant economic impacts. According to Richter et al., (2017) who reviewed the estimated economic impacts of BVDV in 15 endemic countries, the mean direct losses of BVDV infection was approximately US\$ 200 per dairy cow and US\$ 170 per beef cow.

These significant economic losses due to BVDV infection have driven many countries to implement regional or national BVDV control programmes with either voluntary, compulsory, or phased voluntary to compulsory frameworks (Pinior et al., 2017). Several countries in Europe including Sweden, Norway, Finland, and Denmark have already successfully eliminated the disease within a decade using systematic compulsory strategies tailored to their cattle production system. Follow-up studies have economically justified that BVDV elimination programmes in these countries have incurred greater benefits than costs (Richter et al., 2017). Yet, despite the successful examples from Europe, many countries have not achieved elimination potentially due to the complexities of BVDV epidemiology and/or the lack of a systematic and coordinated approach to disease control. As highlighted by Wernike et al., (2017), the voluntary BVDV control programme originally implemented in Germany was not able to successfully eliminate the disease and the authors argued that systematic compulsory strategies should be adopted in countries with high cattle densities and frequent animal movements. Nevertheless, many countries with BVDV are still struggling to convince farmers and stakeholders to support transition from voluntary to compulsory national BVDV control programme despite the well-established returns on investment from implementing coordinated compulsory national strategies (Richter et al., 2017; Wernike et al., 2017).

1.3. BVDV in New Zealand

BVDV has been endemic in New Zealand since at least the 1960s (Salisbury et al., 1961), and almost 50% of both dairy and beef farms in the country have experienced recent BVDV exposure (Gates, Han, et al., 2019). Although it has not been properly estimated, the economic impact of BVDV infection on the New Zealand dairy industry is speculated to be NZ\$ 150 million per year (Anonymous, 2015) and its impact on the beef industry is also believed to be significant.

BVDV control in New Zealand has been based on farmers' voluntary engagement with control measures without any financial support from the government of cattle industries. Over the past decade, some progress has been made with current *ad hoc* voluntary scheme especially in the dairy industry, reducing the national prevalence of BVDV positive dairy farms from approximately 14% to 7% (Gates, Han, et al., 2019; Thobokwe et al., 2004). However, the overall level of farmer compliance to recommended BVDV management strategies in the current scheme is still low, with only ~30% of farmers with recent BVDV exposure conducting a follow-up test to identify persistently infected (PI) animals (Gates, Han, et al., 2019).

1.4. BVDV elimination in New Zealand

Dowdle, (1999) defined the disease control as a "reduction of incidence, prevalence, morbidity or mortality to a locally acceptable level" whereas the elimination as a "reduction to zero incidence in a defined geographical area", and emphasised that both control and elimination are "the result of deliberate efforts" and require "continued intervention measures to maintain the reduction". Therefore, a control or elimination of livestock disease requires a successful implementation of intervention measures, which should be based on the comprehensive understanding of the epidemiology and economics of the disease in the particular population (Petra et al., 2015). In order to eliminate BVDV in New

Zealand, quality data should be collected (1) to quantify demographic structure (e.g. number of herds, number of animals in a herd, and animal movements between farms) and management features (e.g. planned start of mating, and age of weaning), (2) to quantify important BVDV transmission routes between herds, (3) to measure seroconversion rates to predict BVDV transmission dynamics within a herd, and (4) to benchmark the production level of cattle farms depending on BVDV infection status (Perry et al., 2001). BVDV simulation models can then be developed to examine the feasibility of BVDV elimination and to justify the cost-effectiveness of eliminating the disease in the country (E. Brooks-Pollock et al., 2015).

Given the significant economic impacts of BVDV infections, there has been growing interest in evaluating the cost-effectiveness of national BVDV elimination in the New Zealand cattle industries. However, it was recognised that there were also still many questions to be answered about feasibility and the best approach to eliminate BVDV in the New Zealand's pastoral production systems. To help build the capacity to eliminate BVDV from New Zealand, several research projects have been conducted. For example, MSD Animal Health conducted the "Take the BVD Test Challenge" project from September 2015 to December 2016 to collect data about risk factors for BVDV infection on cattle farms across New Zealand. Also, a three-year "BVD Free New Zealand" project was launched in July 2017 to address the knowledge gaps, particularly with regards to the epidemiology of BVDV in the beef industry. Although not designed specifically for controlling BVDV, the national animal identification and tracing (NAIT) legislation also provides support for national disease control efforts by requiring farmers to record all movements of individual cattle as well as details about locations where the animals were moved, which can enhance our understanding about cattle demographic structure in New Zealand. Analysing these data could address the knowledge gaps about the epidemiology and economics of BVDV and its control in the country to enrich our understanding about BVDV elimination in the New Zealand cattle industries.

1.5. Thesis structure

By using the data described above, the aim of this thesis was to address the epidemiological and economic knowledge gaps about BVDV infection in the New Zealand cattle industries to evaluate the feasibility and economic justifiability of implementing national BVDV control programmes. Specifically, in Chapter 2, the key epidemiological features of BVDV were reviewed and the current state of knowledge about BVDV in New Zealand was discussed. In Chapter 3, a risk factor analysis for BVDV infection in New Zealand dairy and beef farms was conducted using data collected for the “Take the BVD Test Challenge” project. This study suggested that local spread of BVDV between cattle herds on neighbouring farms would be the most important route of BVDV transmission in New Zealand. In Chapter 4, the BVDV transmission rate of persistently infected animals on New Zealand beef farms was inferred based on a cohort study from “BVD Free New Zealand” project that measured BVDV seroconversion rates among first-calving heifers from 75 beef breeding farms. Chapter 5 described the development of within-farm BVDV transmission models for a typical New Zealand dairy and beef farms. Different BVDV control measures including annual test and cull of breeding calves, vaccination, and/or double fencing were implemented within the models to evaluate the cost-effectiveness of different BVDV control strategies for individual farms. By incorporating the findings from the previous chapters on an inferred national cattle demographic structure from NAIT data, a national BVDV simulation model was established in Chapter 6 to identify the most cost-effective national control programmes to eliminate BVDV from the New Zealand cattle industries. These studies altogether confirmed that BVDV elimination is both technically feasible and cost-effective for New Zealand pastoral production system.

This thesis was structured as a series of five distinct research papers formatted for publication. At the time of submission of the thesis, Chapter 2, 3, and 4 have been published in peer-reviewed journals.

Chapter 2

Elimination of bovine viral diarrhoea virus in New Zealand: a review of research progress and future directions

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2.1. Abstract

The significant impacts of bovine viral diarrhoea (BVD) on cattle health and production have prompted many countries to embark on national elimination programmes. These programmes typically involve identifying and removing persistently infected (PI) cattle in infected herds and implementing biosecurity measures such as vaccination and pre- or post-movement testing to prevent the virus from spreading to pregnant dams at risk of creating new PI calves. Although the same tools are available to the New Zealand cattle industries, there are still many unanswered questions about how they should be implemented to achieve the greatest benefits to the industries at the lowest cost to individual farms. The epidemiological situation for BVD in New Zealand is unusual due to the extensive pasture-based and seasonal nature of cattle production, the frequent movements of animal between farms, and the potential role of sheep in disease transmission. In this manuscript, we briefly review the key epidemiological features of BVD and the current state of knowledge about BVD in New Zealand. We introduce the ongoing elimination programmes in Europe and highlight the knowledge gaps that we are aiming to address with a new three-year research

programme to design national elimination strategies that could be implemented in New Zealand.

2.2. Introduction

Bovine viral diarrhoea (BVD) is an OIE-listed disease of cattle that is present worldwide, and is well known for its impact on cattle health and production (Anonymous, 2017c). The significant financial impacts of BVD have prompted a number of European countries to invest in either regional or national elimination programmes designed to fit with each country's unique BVD epidemiologic and social context (Houe et al., 2006). After approximately 10 years from implementation, the Scandinavian countries have demonstrated that elimination of the disease is achievable using systematic and coordinated approaches (Ståhl & Alenius, 2012; Wernike et al., 2017).

Recognising the importance of this disease to New Zealand, the National BVD Steering Committee, which has a primary objective to improve awareness of BVD among farmers and veterinarians, was formed in 2005 with representatives from all veterinary groups in the industry (Ellison, 2011). Over the past decade, significant progress, including widespread use of routine BVD testing for service bulls, expansion of BVD diagnostic tests offered by the commercial laboratories, and increased recognition amongst farmers of the important production impacts of BVD, has been made (Stewart, 2013). Based on those achievements, the committee developed a voluntary BVD control framework for the cattle industries despite a lack of monetary subsidies or legislative support from the government.

Although anecdotal reports suggested that the proportion of dairy herds actively infected with BVD virus (i.e. presence of or recent contact with animals that are persistently infected with BVD virus) have decreased from 2010 to 2013 (Stewart et al., 2014), the estimated economic burden of this disease is still in excess of NZ\$ 150 million per year from direct production losses alone (Anonymous, 2015). A lesson from the early phases of BVD

control in Germany was that a voluntary programme has only a limited effect (Wernike et al., 2017), and it has been demonstrated that elimination of BVD in countries with high animal density and movements is not possible without implementing a systematically coordinated scheme (Moennig, Houe, et al., 2005). Therefore, a framework-shift from the current *ad hoc* voluntary situation to a more systematic approach would be required to eventually eliminate BVD in New Zealand.

In order to design a systematic national control programme that achieves the greatest benefits to the industries at the lowest cost to individual farmers, an accurate understanding of the epidemiology, economics, and social motivation for BVD control in New Zealand is necessary. Although a number of BVD research projects have been undertaken in New Zealand, there are still many questions that must be answered to design a cost-effective systematic scheme to control the disease. This article reviews the current state of knowledge about BVD in New Zealand and highlights the knowledge gaps that need to be addressed. We first provide a brief review of the important epidemiological features of BVD, and then describe current knowledge of prevalence, risk factors for transmission, and financial impacts of BVD in New Zealand. BVD control frameworks from other European countries are then discussed along with the validity of BVD diagnostic testing and vaccination in New Zealand. Finally, we report the challenges that must be addressed to design a cost-effective national control programme to eliminate BVD in New Zealand. To help understanding of readers, all financial values are reported in NZ\$ and standardised to the year 2016 (Anonymous, 2016b, 2016a).

2.3. General features of BVD

2.3.1. Biology of BVD

BVD virus is a *Pestivirus* in the *Flaviviridae* family that primarily affects cattle, but can also cause infection in small ruminant species (Evans, Pinior, et al., 2019). Classically,

BVD virus was divided into two genotypes, BVDV-1 and BVDV-2 (Yeşilbağ et al., 2017). With the development of molecular analysis, however, novel strains of BVDV-like pestiviruses, such as "HoBi-like pestivirus" (or BVDV-3), have been identified and studies have shown the worldwide distribution of the strains (Weber et al., 2016). The virus is also classified into two biotypes, cytopathic (cp) or non-cytopathic (ncp), based on the occurrence of apoptosis in infected cells (Grummer et al., 2002).

While the severity of clinical signs varies between different BVD virus strains (i.e. genotype or sub-genotype), most of the acute infections tend to be subclinical (Houe, 2005). After the initial infection, BVD virus is cleared from the host within two to three weeks and this is referred to as transient infection (TI) (Liebler-Tenorio et al., 2004; Müller-Doblies et al., 2004). However, there have been several reports of the virus being maintained in peripheral blood or testicular tissue for prolonged periods (Collins et al., 2009; Voges et al., 1998). Recovered cattle maintain high levels of antibody against that particular BVD virus strain for at least three years (Fredriksen et al., 1999).

Depending on the virulence of BVD virus strain, acutely infected cattle can show clinical signs of fever, diarrhoea, haemorrhage, and immunosuppression which subsequently lead to increased susceptibility to other infectious diseases, and the combination of clinical signs and lethargy often results in reduced milk production, weight loss, and/or reduced growth rates (Baker, 1995; Liebler-Tenorio et al., 2002). BVD virus can also reduce the volume of semen and affect the fertility of female cattle (Kommissrud et al., 1996; McGowan et al., 2003). Reproductive disorders, such as abortions/stillbirths, congenital malformations, or the birth of unthrifty calves, can occur in pregnant cows or heifers depending on the stage of gestation at the time of infection (Moennig & Liess, 1995).

Importantly, if pregnant cows or heifers are infected between approximately 30 to 125 days of gestation and successfully give birth, those calves can be persistently infected (PI) with BVD virus (Grooms, 2004; Moennig & Liess, 1995). Only ncp strains of BVD virus are capable of generating PI animals, as they inhibit the induction of type I interferon in the

fetus (Peterhans & Schweizer, 2013), causing the fetus to become immune-tolerant to the infecting virus strain. Once born, the PI calf will continuously shed large quantities of the BVD virus throughout its life time (Houe & Meyling, 1991). Generally, PI calves are weak, ill-thrifty, and have reduced life expectancy (Houe, 1993), however, some can be clinically normal and survive beyond three years of age (Voges et al., 2006).

Another feature related to the BVD virus biotype is the occurrence of mucosal disease (MD), which is the result of PI animals becoming superinfected with a homologous cp strain (Brownlie et al., 1984). While the recombination of an innate ncp strain with either host RNA or a newly introduced heterologous cp strain can create a homologous cp strain (Goens, 2002), MD can also be developed by a point-mutation of an innate ncp strain itself (Kümmerer et al., 2000). Once this occurs, MD is invariably accompanied by serious clinical manifestations of fever, anorexia, gastrointestinal erosions, severe diarrhoea and eventually death (Houe, 2005). A more detailed review of the pathogenesis of BVD/MD has been provided by Lanyon et al., (2014).

2.3.2. Transmission

Direct contact with PI animals is the most important route of BVD transmission in most cattle herds (Houe, 2005; Moen et al., 2005). Introduction of a PI animal can result in devastating consequences in a naïve herd as the virus can spread rapidly resulting in up to 97% of susceptible cows seroconverting (Houe & Meyling, 1991). BVD can also be transmitted over the fence between neighbouring farms through direct nose-to-nose contact. In Denmark, 67 cases of BVD virus introduction in dairy herds were investigated; the most common transmission pathway was contact with PI animals in a neighbouring herd (36%), followed by the purchase of PI animal (28%) (Bitsch et al., 2000). The purchase of Trojan dam (i.e. a non-PI pregnant cow/heifer carrying a PI fetus) is another important BVDV transmission pathway between farms, and the movement of Trojan cows has been estimated

to contribute less 10% of BVDV incidences under endemic situation (Graham et al., 2014; Qi et al., 2019; Reardon et al., 2018). Grazing on common pastures with stock from other properties also increases the risk of contact with PI animals. In Norway, the odds of having active BVD virus infection was 5.1 (95% confidence interval: 1.97 ~ 13.19) times higher in herds that grazed their heifers on shared pasture (Valle et al., 1999). Compared to PI animals, TI animals are considered to have a limited impact on BVD transmission, as several studies found that TI animals cannot propagate BVD in a herd, presumably due to the short duration and significantly smaller amounts of virus being shed (Niskanen et al., 2002; Niskanen & Lindberg, 2003). BVD transmission via other indirect routes is possible through contaminated biological materials (e.g. vaccines), embryos, or semen (Bitsch et al., 2000; Kommisrud et al., 1996; Niskanen & Lindberg, 2003). Likewise, contaminated vehicles or equipment (e.g. needles, gloves or boots) may introduce the virus (Valle et al., 1999).

2.3.3. Diagnosis

A variety of BVD diagnostic tests detecting either evidence of previous exposure (i.e. antibodies against BVD) or active infection (i.e. BVD virus antigen or viral RNA), are available (Lanyon et al., 2014). Depending on the purpose of the test, either individual samples (e.g. identifying a PI animal) or pooled samples (e.g. assessing a herd's BVD virus infection status) can be tested. Currently, BVD antibody and antigen ELISA, and RT-PCR are commonly used in New Zealand (Bradstock et al., 2018), and the performance of these tests is generally reliable (Dubovi, 2013).

BVD antibody ELISA tests using serum or milk samples can detect previous exposure to BVD virus in individual animals or groups of animals. BVD antibody ELISA test is often used on pooled samples as a screening test to determine herd's immune status against BVD virus (Houe et al., 2006). Also, a highly positive bulk tank milk (BTM) antibody

ELISA, or a high proportion of seropositive individual serum samples from young stock can be considered as an indicator of active BVD virus infection (Beaudeau, Belloc, Seegers, Assié, Pourquier, et al., 2001; Houe, 1994). Although the BVD antibody ELISA test is a cost-effective method to initially assess a herd's immune status, the test cannot be used if target herds are BVD vaccinated, or if calves are young enough to have maternal antibodies. If an initial screening test indicates previous exposure to BVD virus, it is then necessary to conduct further tests to identify any PI animals.

BVD antigen ELISA tests can be applied to milk, serum, or tissue samples to identify PI animals, and has been demonstrated to be a rapid method with high test performance (Kuhne et al., 2005; Saliki et al., 2000). However, serum is an inappropriate sample when testing animals with maternal BVD antibodies as they can interfere with the reaction process of the BVD antigen ELISA test (Fux & Wolf, 2012). The BVD antigen ELISA test also has limited performance if it is used on pooled samples to detect PI animals (Bedekovi et al., 2012; Cleveland et al., 2006).

RT-PCR is known to be highly sensitive to detect viral RNA, at up to 1:125 dilution of serum and 1:600 dilution of milk samples of a PI animal being able to be detected by the test (Bedekovi et al., 2012; Renshaw et al., 2000). However, it is important to recognise that RT-PCR may not differentiate between persistent and transient infection due to its high sensitivity (Torstein Sandvik, 2005). To differentiate between PI and TI animals, one can retest with the antigen ELISA after three weeks to determine if the animal is still viraemic; a TI animal would no longer be positive at the second test (Lanyon et al., 2014). Greater detail about the different types of BVD diagnostic tests and their performance or caveats are thoroughly described by Dubovi, (2013) and Lanyon et al., (2014).

2.3.4. Vaccination

The primary objective of BVD vaccination is to confer immunity to breeding cattle before mating to prevent the creation of PI fetuses (Ridpath, 2013). Although the efficacy of BVD vaccination to prevent fetal infection varies according to the timing of vaccination, type of vaccine (e.g. modified-live or killed), or cross-reactivity between different BVD virus strains (Kelling, 2004), a recent systematic review of 34 previous studies on BVD vaccination reported that vaccines significantly lowered the risk of fetal infection by 85.7% (Newcomer et al., 2015).

BVD vaccination could be beneficial in national control programmes, especially where cattle farms are densely clustered with frequent movements between them (Moennig, Eicken, et al., 2005). BVD vaccination can help prevent the creation of new PIs by keeping a herd resistant to new introductions of BVD virus while clearing existing BVD virus infections (Rypuła et al., 2013). However, vaccination can have some untoward side-effects as the presence of vaccine-induced antibody interferes with the ability to estimate a herd's BVD exposure status using BVD antibody based test (Ridpath, 2013). Also, because of the antigenic diversity of BVD virus, vaccine-induced immunity may not provide sufficient protection against some field BVD virus strains (Brock & Cortese, 2001). Although it is not commercially available anymore, one commercial vaccine was reported to induce a fatal neonatal pancytopenia in calves of vaccinated dams (Jones et al., 2013). The use of modified live vaccines on clinically normal PI animals may result in the occurrence of MD (Ridpath & Bolin, 1995). The risk of BVD vaccines being contaminated with BVD virus also exists (Ridpath, 2013). Most importantly, using a vaccine may give a false impression of safety to farmers, making them less concerned about biosecurity (Lindberg et al., 2006). With or without vaccination, the risk of BVD transmission from other herds still exists if possible transmission pathways are not properly managed, and it has been pointed out that biosecurity is the most important component of BVD control (Ståhl & Alenius, 2012).

Therefore, a vaccination-alone policy should be avoided to achieve successful BVD elimination (Ridpath, 2013).

2.4. BVD in New Zealand

2.4.1. Prevalence of BVD

BVD was first recognised in New Zealand in the 1960s through a case report of mucosal disease-like clinical signs in a herd of cattle (Salisbury et al., 1961). Subsequently, BVD virus was isolated from multiple herds confirming active circulation of the virus within the country (Fastier & Hansen, 1966; Jolly et al., 1967). More recent studies have provided updated estimates of prevalence, further confirming that BVD is endemic and widespread in both the dairy and beef sectors. Only one genotype, a sub-genotype of BVDV-1a, has been reported in New Zealand (Packianathan et al., 2017; Vilček et al., 2001), however, no studies investigating the genetic diversity of the virus in New Zealand have been published during the last two decades (Vilček et al., 1998).

2.4.1.1. Prevalence in the dairy sector

It has been reported that a large proportion of New Zealand cattle herds are affected by BVD. A study that reported BVD antibody ELISA results on BTM from 350 dairy herds across New Zealand showed that 92.9% of the dairy herds had evidence of exposure to BVD virus (Voges, 2008). Assuming that a higher reaction to the ELISA indicates higher seroprevalence which could be driven by the presence of PI animals, the author concluded that 37.5% of the dairy herds were actively infected with BVD virus. This study may have overestimated the proportion of herds with PI animals since the study was not based on any confirmation test on the presence of BVD virus. More recently, Weir et al., (2016) conducted a similar study on BTM of 385 dairy herds, and applied Ab ELISA and RT-PCR

test to confirm the exposure to BVDV and presence of PI animals, respectively. With some variation by year, the authors found that 63.1 ~ 70.1% of the dairy herds had evidence of exposure to BVD virus. Based on RT-PCR, it was suggested that 8.6% of North Island herds and 32.0% of South Island herds contained at least one PI animal. As pointed out by the authors, one limitation of this study was that it was based on herds conveniently selected from only some regions of the country, so the results may not be representative of other parts of New Zealand.

2.4.1.2. Prevalence in the beef sector

Compared to the dairy sector, there have been fewer studies estimating the prevalence of BVD virus infection in New Zealand beef herds due to the greater logistical challenges of conducting diagnostic tests in extensively managed herds (Sanhueza et al., 2013). Heuer et al., (2008) reported that 64.9% (61 out of 94) of beef herds in New Zealand had evidence of previous exposure to BVD virus. Based on the seropositivity of young stock, the authors implied that 46.8 ~ 50.7% of New Zealand beef herds were actively infected with BVD virus. A similar estimate was reported by a recent study using the BVD antibody ELISA test of individual serum samples. Assuming that more than 20% seropositivity in a sampled herd indicates active BVD virus infection, W. Cuttance & Cuttance, (2014) estimated that the proportion of beef herds in the North Island with at least one PI animal was 58.1% (25/43). However, the estimated prevalence of actively infected beef herds may be inaccurate because (1) none of the previous studies confirmed the presence of actively circulating virus in the herds with evidence of exposure, and (2) only a small number of beef farms participated in each study and these were either willing to test for BVD or had a high index of suspicion for BVD, which makes these herds potentially non-representative of the general population. Further studies including larger numbers of beef farms across New Zealand and confirming the presence of BVD virus are required to accurately estimate the prevalence.

2.4.1.3. Within-herd prevalence of PI animals

The prevalence of PI animals within BVD virus infected herds in New Zealand is reported to be less than 2% for both dairy and beef herds (W. Cuttance & Cuttance, 2014; Voges et al., 2006), similar to other countries (Fulton et al., 2009; Rüfenacht et al., 2000). Even though PI animals often have a short life expectancy, there have been several reports that PI cows can survive for at least three years under New Zealand farming conditions (Hill, Reichel, et al., 2010; Voges et al., 2006).

2.4.2. Risk factors for transmission

Several studies have been conducted to identify the risk factors for BVD transmission in New Zealand. Previous studies suggested that factors, such as the purchase of cattle, grazing dams away from the home farm during early pregnancy, or contact with neighbouring cattle, are significantly associated with active BVD virus infection (W. Cuttance & Cuttance, 2014; Weir & Heuer, 2009). There is, however, conflicting evidence that herd size is a risk factor with several studies suggesting that herd size was not a risk factor for BVD virus infection in New Zealand (Thobokwe et al., 2004; Voges, 2008) and another study showing significantly increased risk with increasing herd size (Weir & Heuer, 2009). Estimates of regional risk of BVD virus infection vary as well. Significant geographical patterns in disease prevalence were observed for New Zealand dairy herds, indicating that local spread and local trade networks may influence BVD transmission (Voges, 2008). In contrast for beef herds, W. Cuttance & Cuttance, (2014) found that there was no significant difference in herd level prevalence of active BVD virus infections in beef herds across different regions. This discrepancy around the importance of region as a risk factor could originate from the inherent differences in the management systems between dairy and beef herds (Gates et al., 2014).

2.4.3. Financial losses due to BVD

As well as causing direct production losses in infected farms, the cost of preventive measures on naïve farms, such as vaccination or diagnostic testing, significantly increases the economic burden of BVD to the cattle industries (Richter et al., 2017). Many studies have been conducted internationally to quantify the financial losses due to BVD infection. A recent systematic review estimated that the direct losses due to BVD virus infection ranged from NZ\$ 3.45 to NZ\$ 988.04 per animal (Richter et al., 2017). It should be noted that the estimated losses greatly vary between studies as they applied different estimation approaches (e.g. type of quantitative model, target population, study design, stock and production values, or disease effects contributing to economic losses) as well as different epidemiological settings (e.g. virulence of virus strain, duration of BVD, within- or between-herd prevalence, or herd production level). Nevertheless, it is well recognised that BVD causes substantial economic losses to cattle industries (Houe, 1999).

Given the widespread prevalence of BVD throughout New Zealand, it is likely that BVD has a significant impact on the national economy. Accordingly, the economic impact of BVD in New Zealand cattle herds has been estimated. Several studies have reported significant growth retardation and reduced milk production in New Zealand cattle herds containing PI animals (Compton & McDougall, 2005; Hill, Reichel, et al., 2010; Voges, 2008). Compton et al., (2006) suggested that infertility, abortion, and reduced milk production induced by BVD virus infection caused an annual loss of NZ\$ 109.30 per cow in a dairy herd with PI animals. Assuming 304 cows per herd and 17% of national herds carrying PI animals, the authors estimated the economic impact of BVD to be NZ\$ 71.5 million in the dairy industry alone. Using modelling approaches, Heuer et al., (2007) also analysed the financial loss in the New Zealand dairy industry due to BVD virus infection; assuming an average dairy herd size of 215 with 14.6% of those herds having at least one PI animal, the authors estimated losses to the dairy industry of NZ\$ 52.8 million per year. These studies likely underestimate the true cost of disease as indirect effects, such as

diagnostic testing, vaccination, and increased susceptibility to other infectious diseases, were not considered. Another study calculated that acute infection caused NZ\$ 91.10 per adult dairy cow due to discarded milk, reduced conception and pregnancy rates, and extra veterinary costs (Weir et al., 2016). In the New Zealand beef industry, one study suggested that BVD virus infection resulted in a 5% decrease in pregnancy rate (Heuer et al., 2008). However, no studies have been conducted to estimate the financial cost of BVD virus infection in the beef sector.

2.5. Control of BVD

2.5.1. European programmes

The three critical principles of BVD control in cattle herds are (1) the timely identification and elimination of PI animals to break the within-herd transmission cycle, (2) improved herd biosecurity to prevent BVD virus introduction or re-introduction from outside sources, and (3) continuous surveillance to early detect the re-introduction of BVDV to farms that achieved freedom (Lindberg & Houe, 2005). The latter is why many European countries (e.g. Sweden, Norway, Finland, Denmark, Austria, Switzerland, Germany, the Netherland, Belgium, Scotland, and Ireland) have implemented regional or national systematic BVD control programmes (Moennig & Becher, 2018; Richter et al., 2017). The frameworks used to eliminate BVD in those countries can be generally categorised into three different groups.

2.5.1.1. Scandinavian framework

This framework was adopted in Sweden, Norway, Denmark, and Austria (Houe et al., 2006). The programme has three stages; (1) An initial BVD antibody screening test is conducted for each herd using bulk tank milk (BTM). If positive, or in the case of beef

herds, targeted BVD antibody testing using individual serum from young stock is performed to confirm recent exposure to BVD virus. (2) All individual animals and calves born in the following year in BVD-positive herds are then tested for virus to identify and eliminate PI animals. (3) Once PI animals have been eliminated, the herd's BVD virus infection status is continually monitored by targeting representative young stock, such as weaned calves or replacement heifers. During the programme, movement restrictions for PI animals or animals from infected herds can be imposed to prevent onward transmission.

2.5.1.2. Swiss framework

In the initial phase, all animals were individually tested for BVD virus using an ear notch sample and/or serum sample. In the calf phase, after removal of existing PI animals, any new born calves are screened with either the RT-PCR or the BVD antigen ELISA in case there were Trojan cows at the time of initial screening. For herds with calves positive on this screening test, movement of all pregnant females in the herd is restricted until all the calves are born. In the monitoring phase, after the removal of all PI animals, each herd's BVD virus infection status is monitored with a BVD antibody test on BTM or serum samples. Although the original plan was to end the calf phase by 2009, the phase was continued until 2012 and an antibody-based monitoring programme has been in place since 2013 (Presi & Heim, 2010; Stalder et al., 2016).

2.5.1.3. German framework

This programme is similar to the Swiss framework in that it uses a BVD antigen test for screening. However, rather than testing all cattle on the farm, only the new born calves are tested within 6 months of birth. Another distinct feature is that, unlike other frameworks, voluntary BVD vaccination is allowed. Ireland has also adopted this framework to control BVD (Graham et al., 2014; Wernike et al., 2017).

Based on these frameworks, some modification can be applied depending on the country's epidemiological circumstances. For example, the initial screening process in Norway had additional BVD antibody tests on BTM from primiparous cows and pooled serum from young animals (Løken & Nyberg, 2013). The BVD programme in Belgium is similar to the German framework, but extends the target animals to the dams of the tested calves for initial screening (Steven, 2015). In Scotland, where BVD vaccination is allowed, a BVD antibody test using serum samples from calves is conducted on negative herds for monitoring while virus detection using serum or tissue samples from either all the calves or all the animals in the herd is performed on non-negative herds (Anonymous, 2016c).

2.5.2. Validity of diagnostic tests in New Zealand

Several studies have validated the diagnostic tests used in New Zealand. Thobokwe et al., (2004) compared the test results of the p80 antibody blocking ELISA (IDEXX BVDV Total Ab Test; IDEXX Laboratories Inc., Westbrook MA, USA) on different samples of different animal groups from 50 dairy farms in the North Island. The authors analysed the results of the ELISA using BTM of milking cows and serum samples from 15 heifers on each farm, and showed that there was a significant correlation between the results. Assuming that (1) the presence of seropositive heifers indicates active infection of the herd with BVD virus, and (2) heifers and milking cows are managed separately in New Zealand, this study implied that a high level of BVD antibody detected in an ELISA on BTM can be evidence of active infection of BVD virus in a dairy farm. Validation of the BVD antibody test at individual cow level was conducted by Weir et al., (2013). Based on anecdotal reports that the performance of the indirect antibody ELISA test (IDEXX BVDV Total Ab Test; IDEXX Laboratories Inc., Westbrook MA, USA) on individual milk samples was suboptimal, the authors analysed the test results of the ELISA using milk and serum samples from 90 cows. This study suggested that the cut-off value for the test using milk samples should be lowered to correctly identify exposed individual animals. With a

recalibrated cut-off value, the relative sensitivity and specificity for the milk sample compared to the same test on serum samples were estimated as 0.99 and 1.00, respectively.

Hill et al., (2007) validated the BVD antigen ELISA test (IDEXX Ag-ELISA; IDEXX Scandinavia, Österbybruk, Sweden) using skin and serum samples from 276 animals. For 30 PI animals confirmed by RT-PCR test, ELISA test results using both samples showed complete agreement as well as with the RT-PCR test. This study indicated that using a skin sample for the BVD antigen ELISA test could be a valid and cheaper approach to accurately detect PI calves in New Zealand. RT-PCR on BTM has also been validated to be sensitive to detect PI animals in New Zealand dairy herds. Hill et al., (2010b) demonstrated that RT-PCR on BTM detected 2 PI animals among a herd of 800 cows. However, the test failed to detect 1 PI animal among 800 cows, implying that RT-PCR on BTM was able to detect 1 PI animal among up to 400 negative cows in New Zealand. Based on the result, the authors argued that RT-PCR on BTM was a reliable screening test in New Zealand dairy herds.

2.5.3. Validity of BVD vaccination in New Zealand

In New Zealand, McArthur, (2004) assessed the efficacy of fetal protection of a trivalent vaccine and reported that two out of eight vaccinated cows produced PI calves and a further cow aborted, resulting in a vaccine efficacy of 62.5%. A more recent study assessed the efficacy of two commercially available BVD vaccines (BVDV-1a and BVDV-1c vaccines) in New Zealand (Packianathan et al., 2017). Twenty-five heifers were vaccinated and boosted against BVD virus, and the heifers were challenged with BVDV-1a after 6 months. Of the 22 fetuses where the authors could retrieve fetal tissue, 15 were successfully protected (e.g. 5 out of 8 for the type 1a vaccine, 10 out of 14 for the type 1c vaccine) giving 62.5% (95% confidence interval: 29.0% ~ 96.0%) and 71.4% (95% confidence interval: 47.8% ~ 95.1%) of fetal protection efficacy against BVDV-1a and BVDV-1c strain, respectively.

2.5.4. Financial benefit of BVD control programmes

Many studies have been conducted worldwide to assess the financial cost and benefit of implementing BVD control programmes. It should be noted that the estimated cost and benefit of BVD control programmes varies greatly between studies depending on the assumptions around the target populations, economic parameters, prevalence of BVD, timescale, or type of control measures. However, a crucial point is that most of the previous studies on the financial assessment of BVD prevention programmes worldwide concluded that the implementation of a BVD control programme was economically beneficial (Pinior et al., 2017).

The applicability of some control schemes has been financially evaluated for New Zealand dairy farms. Reichel et al., (2008) conducted a cost-benefit analysis of three different BVD control programmes (i.e. vaccination, test-and-cull, and increasing biosecurity) with varying efficacy of protection, and found that increasing the biosecurity (specifically screening purchased cattle for BVD and enhancing farm boundary biosecurity) resulted in maximum benefit. Another study with a similar aim was conducted testing more options for control schemes and concluded that the combination of double fencing, screening of in-coming animals, and PI elimination was the most cost-effective option (Weir et al., 2014). In both studies, the authors concluded that implementing any form of control measure for BVD benefited farm profitability significantly more than doing nothing. No studies to date have been conducted to estimate the financial benefit of BVD control measures in the New Zealand beef industry. It will be crucial to generate good estimates for both the beef and dairy sectors, given the high degree of contact between them (Anonymous, 2017b).

2.6. Discussion

Today, many European countries are running BVD control programmes regionally or nationally, and some of these countries have virtually eliminated the disease (Richter et al., 2017). Although running a BVD control programme costs significant money, time, and effort, there have been multiple simulation studies demonstrating that the benefits of BVD elimination far out-weigh the cost of the disease control programme (Pinior et al., 2017). An anecdotal report estimated that at least NZ\$ 150 million per year is lost due to BVD in the New Zealand cattle industries (Anonymous, 2015). Considering its major impact on the industries, the general consensus on implementing a systematically coordinated national BVD control programme should be built to increase the productivity of cattle farms. Successful control of BVD will not only reduce the economic burden of the disease, but also increase the animal welfare by preventing the stress and pain experienced by animals infected with BVD virus.

In the past decade, the science around BVD in New Zealand has advanced significantly and we now have the technical capability to eliminate the disease. However, there are still some questions that should be answered before we can determine the most cost-effective tools to employ for a national control programme.

Firstly, most studies on BVD in New Zealand have been conducted in the dairy industry and there is limited knowledge about the disease in the beef sector. This is mainly because beef herds in New Zealand are extensively grazed with few mustering events during the production season (Geenty & Morris, 2017), which makes it difficult for researchers to collect accurate epidemiological and economic data, such as live weight gain for susceptible, TI, or PI animals (Sanhueza et al., 2013). Another epidemiological challenge associated with extensive beef herds is that calves typically stay with their dams from birth to weaning at approximately 6 months of age, which means there are limited opportunities to identify and remove PI calves before they can expose susceptible dams during the next mating period. This is particularly a concern if the contact rate between cattle in extensive

beef herds is not high enough to generate sufficient immunity to BVD before the start of mating if there are PI calves in contact with the herd of breeding cows. Given that there are twice as many beef herds (3.5 million animals from 31,000 herds) as dairy herds (6.5 million animals from 14,000 herds) in New Zealand (Anonymous, 2017a; Jewell et al., 2016), knowledge of BVD epidemiology and its economic impact in the beef sector would be a significant requirement for understanding BVD in the New Zealand cattle industries.

Secondly, to the best of our knowledge, the only analysis on the genotype of endemic BVD virus strains in New Zealand was published in 1998 (Vilček et al., 1998), and there is limited information about the genotypes of currently circulating strains. Given the variety of virulence or antigenic diversity between BVD virus strains, knowledge around genotypes of current BVD viruses should be updated. Molecular epidemiology has also proven to be a useful tool to increase our understanding of BVD transmission and contact patterns in countries that maintain molecular databanks of prevalent BVD virus strains (Ståhl et al., 2005). For example, researchers can use the genetic similarity of BVD virus sequences isolated from different farms to make inferences about the timing and direction of transmission (Booth et al., 2013). By investigating these epidemiological links, we can make better inferences about the most important routes of BVD transmission between cattle farms in the industries as well as better biosecurity recommendations to individual farms to help them target the specific pathways that are leading to BVD virus re-introduction.

Although BVD is typically a cattle disease, it has been demonstrated that other small ruminants, like goat or sheep, can be infected with BVD virus (Passler & Walz, 2010). Even though sheep are considered spill-over hosts of BVD virus infection, there was a report of a PI ram being identified as the source of ongoing BVD virus circulation in a Danish cattle herd for several years (Bitsch et al., 2000). In New Zealand, one recent case report showed that BVD PI lambs in a Wairarapa sheep flock were generated by contact with PI heifers (King, 2014). Considering that co-grazing with sheep, either concurrently or successively, on the same pasture is common in the New Zealand beef industry, future

studies should determine if there is the potential for sheep to be contributing to BVD transmission in New Zealand beef herds.

BVD transmission between herds via animal movements is another fundamental factor that must be understood when establishing a BVD control scheme for New Zealand. In 2016, more than 2.6 million cattle were reported to move between farms in New Zealand (Anonymous, 2017b). This is a significant amount of movement considering that the entire cattle population is estimated to be 10 million (Anonymous, 2017a). Animal movement or trade patterns between New Zealand cattle farms have been researched in relation to other diseases, such as foot-and-mouth disease or bovine tuberculosis (Hidano et al., 2016; Sanson, 2005), but the impact of cattle movement on BVD spread has never been studied. Given that the most important risk factor for BVD virus infection is direct contact with PI animals, the epidemiological role of a PI animal or a Trojan cow introduced via animal movements on national BVD spread needs to be investigated.

New Zealand farmers' opinion and their willingness to adopt BVD control measures should be investigated as well. It has been shown that providing farmers with information about the economic loss caused by a disease is not sufficient to make them implement any control measures (van Asseldonk et al., 2010). In Ireland, nearly 30% of dairy farmers that were confirmed as actively infected with BVD virus retained identified PI animals on the farm even though some financial compensation was available for culling (Graham et al., 2014). A recent study of English farmers' behaviour around biosecurity to prevent the introduction of BVD showed that the farmer's affiliation to the voluntary BVD control scheme was not associated with either their knowledge or level of concern about the disease (Azbel-Jackson et al., 2018). International studies have shown that a variety of factors, such as herd size, market-driven economic premiums, and social and cultural significance (e.g. being recognised as "good farmer" by others), can drive farmers to implement biosecurity measures (Garforth et al., 2013; McAloon et al., 2017; Nöremark et al., 2010). Therefore, studies on the social science that elucidate what motivates New Zealand farmers to join and

comply with control schemes would highlight ways to increase participation among individual farmers who currently have a low commitment to BVD control.

A three-year BVD research project (www.bvdfree.org.nz) sponsored by the Sustainable Farming Fund, AGMARDT, and industry was launched in July 2017 to help build the business case for BVD elimination in New Zealand. This involves field studies and national surveys in the first two years to collect better data on general herd management, BVD epidemiology (e.g. seroconversion rates, prevalence of PI calves, or decay rate of maternal antibody on calves) and its economic impacts, efficacy of control measures (e.g. PI culling or vaccination), and farmers' opinions towards BVD control in New Zealand. The role of sheep in BVD transmission will also be preliminarily investigated, and animal movement patterns between cattle farms and genotypes of current BVD virus strains will be analysed to identify important transmission routes between herds. Data gathered through the first two years will be used to build a national disease simulation model to evaluate the long-term cost-effectiveness of different national elimination options in New Zealand. We expect that most of the epidemiological and economic questions will be answered through the BVD research project, and it should lead to the most cost-effective national BVD control strategy that is relevant to the New Zealand farming context.

2.7. Acknowledgements

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Chapter 3

Using Bayesian network modelling to untangle farm management risk factors for bovine viral diarrhoea virus infection

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3.1. Abstract

Understanding risk factors for bovine viral diarrhoea (BVD) transmission is important for planning national disease control programmes. However, traditional statistical approaches may miss important features of BVD epidemiology due to the highly correlated nature of many farm-level risk factors. In this cross-sectional study, we used data collected from 304 cattle herds in New Zealand during 2015/2016 to compare the results from multivariable logistic regression with Bayesian network (BN) analysis. Blood samples from 15 heifers from each farm were pooled and analysed with an antibody ELISA test to classify BVD virus exposure status. Farmers were surveyed about their general management practices, knowledge about BVD, and risk factors for disease transmission, including onto- and off-farm movements, within- and between-farm contacts, and whether they implemented BVD control measures for their service bulls. Multiple imputation was used to infer missing values in the dataset prior to statistical analysis. The results showed that 57/116 (49.1%) beef farms and 95/188 (50.5%) dairy farms were likely to be actively

infected with BVD virus. Almost 60% of farms had movements of heifers/cows onto the premises and 13.8% of farmers reported contact with cattle from other farms. The results of the multivariable logistic regression showed that farms where heifers/cows had been moved onto the premises during all or most of the past five years were at higher risk of being BVD seropositive than farms without those movements (odds ratio: 2.21, 95% confidence interval: 1.29 ~ 4.24). Farms where cattle had occasional or rare contacts with cattle on other farms were also at increased risk compared with farms without any animal contacts between farms (Odds ratio: 2.63, 95% confidence interval: 1.33 ~ 5.41) although this association was not frequency-dependent. Only close animal contacts between farms was directly associated with BVD status in the BN model, however, this approach further untangled other complex associations between correlated management factors, and provided additional important insights into BVD epidemiology. Compared to other countries with intensive production systems, over the fence contact appeared to play a more important role in New Zealand pastoral-based production systems and should be considered when developing strategies for a national BVD control programme.

3.2. Introduction

Bovine viral diarrhoea (BVD) is an OIE-listed cattle disease that exists worldwide and is recognised for its significant impact on cattle health, welfare and production (Anonymous, 2017c). A distinct characteristic of BVD epidemiology is that, if dams are infected in early-mid gestation before the fetus has developed a competent immune system, the fetus will become persistently infected (PI) with BVD virus and the resulting calf will shed large quantities of the virus throughout its life-time (Fray et al., 2000). Control schemes for BVD in many countries are therefore focused on identifying and eliminating existing PI animals in a timely fashion and preventing the creation of new PI animals through vaccination and improved herd biosecurity (Lindberg & Houe, 2005).

Many studies have used standard multivariable regression to explore risk factors for BVD with the aim of identifying important transmission pathways to improve herd biosecurity plans and national disease control programmes. In particular, the purchase of heifers or cows, mean number and distance of neighbouring farms, sharing a common pasture, or over-the-fence contacts between herds have been shown to significantly increase the risk of BVD transmission (Ersbøll & Stryhn, 2000; Obritzhauser et al., 2005; Valle et al., 1999), and the sharing of farming instruments or contaminated biologicals have also been identified as possible transmission routes (Lindberg & Houe, 2005). These risk factors for BVD transmission are generally highly correlated farm management factors (Gates et al., 2013; Lewis & McCormick, 2012) and using multivariable approaches that only consider the association of variables with one specific outcome of interest may restrict our understanding of how general farm management practices influence BVD epidemiology.

Considering the correlated nature of farm management factors, a multivariate analysis capable of examining complex inter-relationships among all variables of interest may provide a better understanding of BVD epidemiology, which is crucial for better recommendations for BVD control. There are many multivariate techniques, such as factor analysis or principal component analysis, to explore associations between variables, and Bayesian network (BN) analysis is one such multivariate approach that illustrates associations between variables as a directed acyclic graph (DAG) by estimating joint probabilities (Nagarajan et al., 2013). BN modelling empirically discovers an optimal DAG, or network structure, that best represents the associations between variables in a given dataset and can be used to distinguish between indirect and direct associations with the outcome of interest. Although it may require adequate sample size (Zuk et al., 2006), it is generally accepted that BN is not restrained by small sample size (Uusitalo, 2007), meaning that many variables can be considered simultaneously with a restricted number of observations. BN analysis has been increasingly applied in the field of veterinary

epidemiology to analyse complex relationships among variables (Firestone et al., 2014; Pittavino et al., 2017; Schemann et al., 2013) including one pioneering study to holistically analyse BVD-related data (Lewis et al., 2011).

BVD has been endemic in New Zealand since at least the 1960s (Salisbury et al., 1961) and its control currently depends on farmer's voluntary uptake of recommended control measures such as testing and vaccinating bulls and other replacement breeding stock (Han, Weir, et al., 2018). As a part of ongoing efforts to evaluate the cost-efficiency of implementing a coordinated national BVD control programme in New Zealand, the importance of different farm management risk factors of BVD transmission should be evaluated in New Zealand's pastoral farming system. To date, only a limited number of studies have been conducted to identify risk factors in New Zealand (W. Cuttance & Cuttance, 2014; Weir et al., 2016). To address this knowledge gap, MSD Animal Health conducted a "Take the BVD Test Challenge" project in 2015/2016 that surveyed farmers about their current BVD status and management practices. In this study, we analysed the data (1) to investigate risk factors for BVD under New Zealand pastoral farming conditions and (2) to compare and contrast the inferences that can be made about BVD epidemiology from the two different statistical modelling approaches.

3.3. Material and methods

We followed the guidelines from the STROBE-VET statement to report this study (Sargeant et al., 2016).

3.3.1. Data collection

The study was based on cross-sectional data from the "Take the BVD Test Challenge" project conducted by MSD Animal Health from September 2015 to December 2016. The aim of the project was to update information about the prevalence of BVD and to collect data about BVD risk factors from cattle farms across New Zealand. The initiative was

advertised to all veterinarians across New Zealand who were asked to collect whole blood samples for BVD testing and administer a BVD risk factor survey whilst visiting an eligible farm for other animal health reasons. The selection criteria were that the farms (1) were not currently implementing a BVD control programme and (2) were not currently vaccinating animals against BVD virus (with the exception of bulls) since vaccination is known to interfere with antibody-based screening tests for BVD. The source population therefore was New Zealand cattle farms that utilised veterinary services and were not implementing BVD control measures. No restrictions were set on the total number of farms that could participate.

For each enrolled farm, 15 apparently healthy animals between 1 and 2 years of age were randomly selected from the replacement heifer herd. Serum samples were collected from each animal and shipped to a commercial diagnostic laboratory for analysis using a commercially available BVD antibody ELISA test (IDEXX BVDV Total Ab Test; IDEXX Laboratories Inc., Westbrook MA, USA). Farms were inferred as being actively infected with BVD virus if the S/P ratio for the pooled sample was greater than the validated cut-off value of 0.75 (Hill, McCoy, et al., 2010). The laboratory testing fees were provided at no cost to farmers as an incentive to participate.

During the sampling visit, a survey consisting of seven open and 33 closed questions was administered to farmers to collect information about general farm management practices. These included questions on the geographic region and production type of farm, number of neighbouring cattle farms, frequency of off-farm movement of the replacement heifers, frequency of close animal contacts within and between farms, and farmers' awareness of important BVD epidemiological concepts including PI animals and Trojan cows (i.e. a pregnant cow carrying PI fetus). Whether farmers shared cattle yards with neighbour(s) or had onto-farm movements of heifers/cows, bulls, weaners/stores, replacement heifers, or carry-overs (i.e. cows that failed to conceive but were kept in the herd for breeding in the subsequent season) during the last five years were also asked. The

questionnaire also included a mating method (e.g. AI only, bull only, AI & bull, and not available) used for the herd of replacement heifers, and three additional questions about BVD prevention measures employed for service bulls (e.g. whether bulls were BVD antigen-tested, certified BVD-free, or vaccinated against BVD) were further asked for those who replied as either bulls only or AI & bull for the mating method. The same set of questions (mating method and three BVD prevention measures employed for service bulls) was also asked about the mating of mixed-age breeding cows. The survey was reviewed by two local veterinarians to provide feedback on the questions and wording, but no pilot study with farmers was conducted prior to the project. A copy of the survey is provided in Appendix 1 (see Section 1).

3.3.2. Data processing

A total of 308 farmers were voluntarily recruited for the project during the study period. Observations from four farms were discarded due to missing survey data or sampling the wrong population of animals. Among the observations from the remaining 304 farms, overall 5.8% of fields in the dataset were missing data and only 154 (50.7%) farmers completed every question on the management survey. To avoid discarding observations with missing data from the statistical models, we imputed missing values using multiple imputation by chained equations (MICE). MICE creates multiple imputed datasets by sequentially regressing missed values of a variable using other variables as explanatory variables in an appropriate model (van Buuren, 2018). In this study, imputation models of predictive mean matching, logistic regression, proportional odds model, and multinomial logistic regression were used for imputing missing values of continuous, binary, ordinal, and multinomial variables, respectively (White et al., 2011). Multiple imputation was conducted in R using “mice” packages (van Buuren & Groothuis-Oudshoorn, 2011), and 1,000 imputed datasets were created to adjust for variation caused by imputation (van Buuren, 2018). Imputed values of BVD prevention measures for the service bulls used with

the replacement heifers and mixed-age breeding cows were restricted to “Not available” when the observation had either “AI only” or “Not available” for the mating method. Similarly, imputed values for those variables were updated to “Unknown” if the imputed values were “Not available” but the mating method specified the use of service bulls.

Following imputation, several variables were re-categorised and/or merged to simplify the dataset. Responses of variables with more than three categories were re-grouped into two or three categories. Given the fact that only farmers who used bulls for the mating method were able to reply three variables of BVD prevention measures applied to service bulls, these four variables for each group (e.g. replacement heifers and mixed-age breeding cows) were merged, creating variables of “Implementing BVD control measures on service bulls” for mating replacement heifers and mixed-age breeding cows with categories of “Bulls not used”, “No”, and “Yes”. Details on re-categorising and the merging process are provided in Appendix 1 (see Section 2).

3.3.3. Descriptive statistics

Before imputing missing values, basic descriptive statistics were generated using the original dataset. Given the missing values, bi-variable descriptive statistics were conducted in a pairwise deletion manner, which is using only available observations of variables for comparison. For bi-variable comparison, statistical methods of Student’s *t*-test or chi-square test were applied. To estimate the sampling coverage, the number of farms recruited per production type was compared with the number of each type of cattle farms in New Zealand based on census data from 2012 (Anonymous, 2017a).

3.3.4. Multivariable logistic regression

A multivariable logistic regression model was developed to identify risk factors associated with active BVD virus infection. The outcome variable was whether the farm was positive or negative for active BVD virus infection as determined by the S/P ratio from pooled antibody testing, and 19 explanatory variables were evaluated as risk factors. In order to adjust variation between imputed datasets, we stacked 1,000 imputed datasets and applied a fixed weight for regression models (Wood et al., 2008). The weight was $(1 - f)/M$ for each observation, where f is the fraction of missing values across all variables in an original dataset (i.e. 0.055) and M is the number of imputed datasets (i.e. 1,000).

We first performed univariable screening to identify variables that were associated with the outcome at a P-value of likelihood ratio test ≤ 0.2 . During the bi-variable comparison using the original dataset, we observed significant correlation between the frequency of close animal contacts between farms and both the number of neighbouring cattle farms and sharing cattle yard with neighbour(s). Because all the three variables were identified as significant in univariable screening, only the frequency of close animal contacts between farms was considered for multivariable logistic regression. Significant variables from univariable screening were entered in the multivariable model followed by a backward stepwise selection method to identify the final multivariable model. Starting with the highest P-value, any variable with a P-value of the log likelihood ratio test > 0.05 was manually eliminated from the model. After each variable was dropped, we assessed changes in the coefficients and standard error of other explanatory variables to identify possible confounders. The criterion of confounding effect was set as 20%, however no confounding effect was observed during the model building process in this study.

As one of our aims was to compare a multivariable logistic regression model with a BN, interactions between explanatory variables were not considered in this multivariable approach. The result of the final model was presented as odds ratios (OR) and 95%

confidence intervals (CI). All the regression modelling was conducted in R (R Core Team, 2016).

3.3.5. BN modelling

To further investigate associations between risk factors for active BVD virus infection and to identify possible risk factors that were not captured by the multivariable approach, we also constructed a BN model. Construction of a BN involves structure discovery and parameter estimation. Structure discovery is the process of identifying an optimal DAG showing best fit to a given dataset. Since we had 1,000 imputed datasets, an optimal DAG was identified in two steps; (1) identify 1,000 network structures using imputed datasets with a random restart greedy hill-climbing search (Bouckaert, 1995), and then (2) generate an optimal DAG as a majority consensus network among the 1,000 structures (Wilson et al., 2013). Given that the direction of the BN does not indicate causality (Lewis & McCormick, 2012), the majority consensus network was set by retaining arcs with a summed number of appearances over both directions being more than 50% of the 1,000 structures (e.g. undirected majority consensus network) (Poon et al., 2007a). Once an optimal DAG was discovered, the DAG was translated into an additive BN, of which a parameterisation process is similar to the multivariable generalized linear models, to estimate the parameter of each arc in forms of OR (Benjamin J. J. McCormick et al., 2017). For the calculation of OR, direction of undirected arcs in the optimal DAG was re-assigned with the one that appeared at a greater frequency in the 1,000 structures (Poon et al., 2007b). Then, the marginal distributions of every parameter on the optimal DAG were calculated for each imputed dataset, resulting in 1,000 marginal distributions of OR for each arc on the DAG. Those distributions were aggregated to estimate median and 95% credible interval (CrI) of each OR, and an arc was treated as significant if the 95% CrI did not overlap with one. Any arc that was not statistically significant was dropped.

For constructing a BN, all categorical variables with three response groups (except variables of implementing BVD control measures on service bulls for mating the replacement heifers and mixed-age breeding cows) were converted into binary variables in such a way to fit responses into either “Never” or “Not-never”. Production type and implementing BVD control measures on service bulls for heifers and mixed-age breeding cows were split into binary dummy variables, each of which corresponded to the responses of the original variable. For simplicity, the region was dropped as there was no plausible explanation for regional differences. Overall, 24 variables were used for BN modelling, and any arc between split variables was banned. Structure discovery was calculated in R using “bnlearn” package (Scutari, 2009) and parameters were estimated in JAGS 4.3.0 (Plummer, 2013).

3.4. Results

3.4.1. Descriptive statistics from the survey responses

The “Take the BVD Test challenge” was completed by a total of 116 beef farmers and 188 dairy farmers during the study period. The studied farms represented 0.5% of beef and 1.5% of dairy farms in New Zealand. Overall, evidence of active BVD virus infection was observed in 152/304 farms (50.0%, CI: 43.5 ~ 54.8%), including 57/116 beef farms (49.1%, CI: 43.5 ~ 54.6%) and 95/188 dairy farms (50.5%, CI: 44.9 ~ 56.2%). As shown in Figure 3.1, active infection with BVD virus was prevalent in cattle farms in all regions of New Zealand, and there were no obvious variations in prevalence by region.

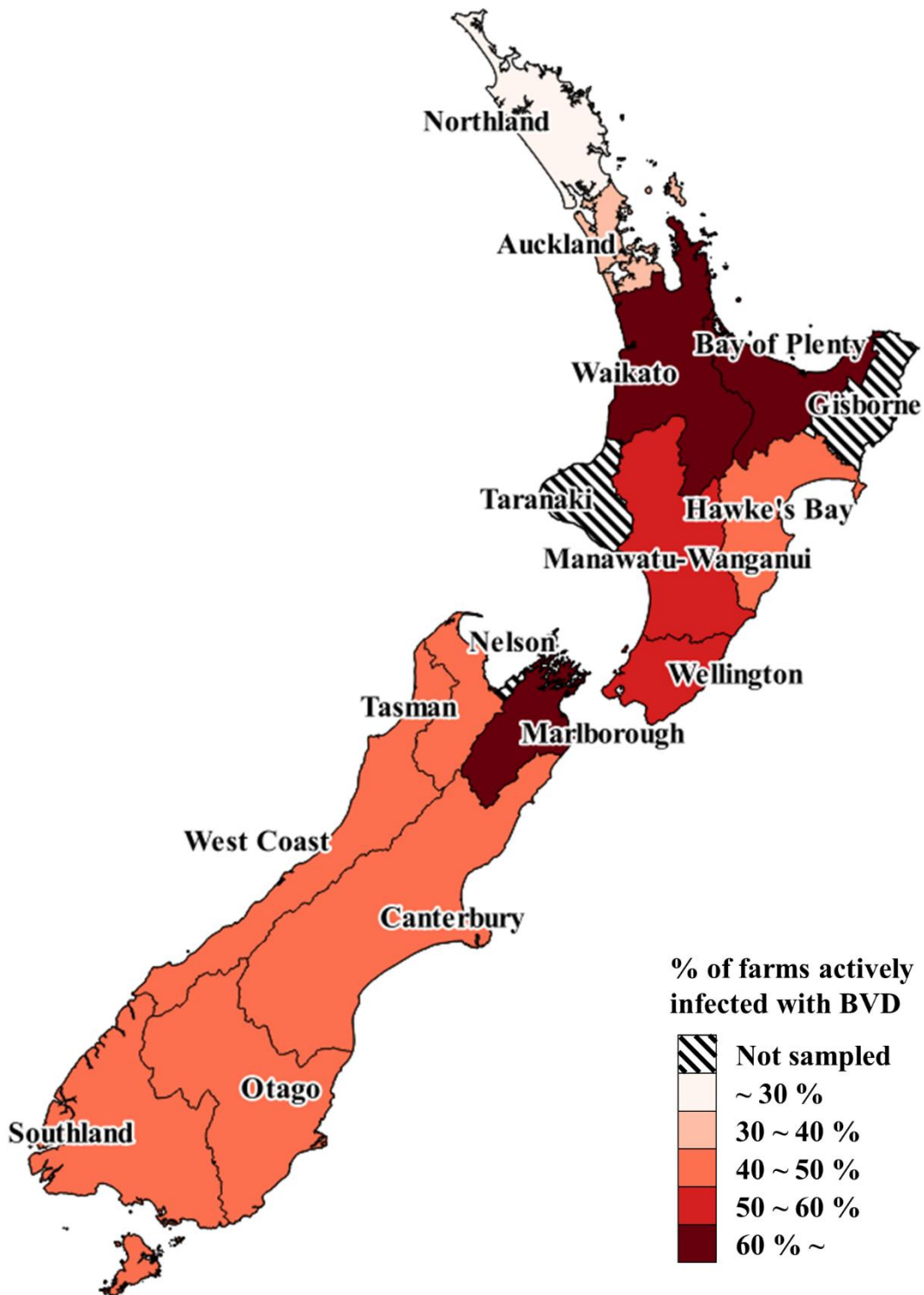


Figure 3. 1. Figure 3.1. Proportion of cattle farms actively infected with BVD. Note that no cattle farms from three regions (Gisborne, Taranaki, and Nelson) were sampled.

Survey responses from the 304 participating farms are summarised in Table 3.1. The overall median number of replacement heifers was 90 animals (IQR 50 to 140) and the number of heifers for dairy farms (median: 103, IQR: 70 ~ 168) were significantly greater (P-value < 0.001) than for beef farm (median: 59, IQR: 30 ~ 100). It was common practice for cattle farms to introduce at least one animal into the herd annually with only 6 out of 285 farmers (2.1%) that provided complete responses about onto-farm movements of cattle reporting that they maintained completely closed herds. Among the 279 farms that maintained open herds, 255 (91.4%) had onto-farm movement of any type of cattle throughout all or most of the last five years. The most common type of cattle brought onto the farm during the last five years was bulls (94.1%), followed by heifers/cows (30.2%), replacement heifers (19.6%), weaners/stores (16.9%), and carry-overs (11.0%). Frequent close animal contact between different management groups within a farm was observed on 72/284 farms (25.4%), while only 40/289 farmers (13.8%) reported to have frequent close animal contacts between animals from different farms. Sharing cattle yards with neighbouring farms was not a common management practice, reported by only 108/297 farmers (36.4%).

Table 3. 1. Description of survey responses and univariable logistic regression of farm management risk factors for active infection with BVD virus from 304 New Zealand cattle herds. Median and interquartile range of number of farms over 1,000 imputed datasets are also presented. Odds ratio with bold and underlined characters shows p-value less than 0.2.

Variable	Level/units	Total number of farms		Number of farms BVD (+) ^{a, b}	OR (95% CI)
		Raw	Imputed ^b		
Region (n= 304)	Northland	9	9	2	Reference
	Auckland	15	15	6	2.33 (0.37 ~ 21.00)
	Waikato & Bay of plenty	46	46	29	5.97 (1.22 ~ 46.54)
	Hawke's bay	24	24	12	3.50 (0.65 ~ 28.93)
	Manawatu-Wanganui & Wellington	18	18	10	4.38 (0.75 ~ 38.00)
	West Coast	13	13	6	3.00 (0.46 ~ 27.74)
	Canterbury	40	40	18	2.86 (0.57 ~ 22.43)
	Otago	40	40	19	3.17 (0.64 ~ 24.80)
	Southland	67	67	31	3.01 (0.64 ~ 22.84)
	Tasman	6	6	3	3.50 (0.36 ~ 43.41)
	Marlborough	26	26	16	5.60 (1.05 ~ 46.26)
Production type (n= 304)	Dairy	188	188	95	Reference
	Beef	116	116	57	1.06 (0.65 ~ 1.72)
Neighbouring farms (n= 274)	Same or less than 2	105	119 (117 ~ 120)	51 (50 ~ 52)	<u>Reference</u>
	More than 2	169	185 (184 ~ 187)	101 (100 ~ 102)	<u>1.61 (1.00 ~ 2.62)</u>
Onto-farm movement of heifers/cows (n= 297)	Never	121	122 (122 ~ 123)	54 (54 ~ 54)	<u>Reference</u>
	Occasionally/rarely	93	94 (94 ~ 95)	43 (43 ~ 43)	<u>1.06 (0.60 ~ 1.86)</u>
	All/most years	83	87 (87 ~ 88)	55 (55 ~ 55)	<u>2.15 (1.20 ~ 3.90)</u>
Onto-farm movement of bulls (n= 302)	Never	24	24 (24 ~ 25)	11 (11 ~ 11)	Reference
	Occasionally/rarely	24	24 (24 ~ 24)	14 (14 ~ 14)	1.66 (0.51 ~ 5.54)

	All/most years	254	256 (255 ~ 256)	127 (127 ~ 127)	1.19 (0.50 ~ 2.91)
Onto-farm movement of weaners/stores (n= 294)	Never	193	198 (197 ~ 199)	95 (94 ~ 95)	Reference
	Occasionally/rarely	56	59 (58 ~ 60)	33 (33 ~ 33)	1.38 (0.75 ~ 2.55)
	All/most years	45	47 (46 ~ 47)	24 (24 ~ 25)	1.19 (0.61 ~ 2.33)
Onto-farm movement of replacement heifers (n= 291)	Never	209	212 (212 ~ 213)	101 (101 ~ 102)	Reference
	Occasionally/rarely	29	30 (30 ~ 31)	18 (18 ~ 19)	1.67 (0.75 ~ 3.84)
	All/most years	53	61 (60 ~ 62)	32 (32 ~ 33)	1.24 (0.68 ~ 2.25)
Onto-farm movement of carry-overs (n= 289)	Never	230	236 (235 ~ 237)	114 (113 ~ 114)	Reference
	Occasionally/rarely	29	31 (30 ~ 32)	15 (14 ~ 15)	0.95 (0.43 ~ 2.08)
	All/most years	30	37 (36 ~ 38)	24 (23 ~ 24)	1.88 (0.91 ~ 4.07)
Sharing cattle yards (n= 297)	Never	189	192 (192 ~ 193)	90 (89 ~ 90)	<u>Reference</u>
	Occasionally/rarely	82	84 (84 ~ 85)	51 (50 ~ 51)	<u>1.74 (1.01 ~ 3.00)</u>
	All/most years	26	27 (26 ~ 28)	11 (11 ~ 12)	<u>0.83 (0.35 ~ 1.93)</u>
Number of replacement heifers (n= 270)	(Continuous)	-	-	-	1.00 (1.00 ~ 1.00)
Off-farm movement of heifer herd (n= 298)	No	152	155 (154 ~ 156)	76 (75 ~ 76)	Reference
	Yes	146	149 (148 ~ 150)	76 (76 ~ 77)	1.10 (0.69 ~ 1.76)
BVD control measures on service bulls for heifer herd (n= 259)	Bulls not used	83	87 (86 ~ 89)	43 (42 ~ 44)	Reference
	No	42	52 (51 ~ 54)	30 (29 ~ 31)	1.36 (0.67 ~ 2.81)
	Yes	134	164 (162 ~ 165)	78 (78 ~ 79)	0.93 (0.54 ~ 1.60)
Close animal contacts within a farm (n= 284)	Never	46	50 (49 ~ 51)	20 (19 ~ 21)	Reference
	Occasional/rarely	166	176 (174 ~ 177)	92 (90 ~ 93)	1.59 (0.82 ~ 3.13)
	Often	72	79 (78 ~ 80)	40 (39 ~ 41)	1.51 (0.72 ~ 3.23)
Close animal contacts between farms (n= 289)	Never	43	47 (46 ~ 49)	16 (15 ~ 16)	<u>Reference</u>
	Occasional/rarely	206	215 (214 ~ 216)	119 (118 ~ 119)	<u>2.44 (1.24 ~ 4.97)</u>
	Often	40	42 (41 ~ 42)	17 (17 ~ 18)	<u>1.43 (0.58 ~ 3.55)</u>
Off-farm movement of mixed-age cows (n= 272)	No	171	187 (186 ~ 189)	94 (93 ~ 95)	Reference

	Yes	101	117 (115 ~ 118)	58 (57 ~ 59)	0.98 (0.60 ~ 1.59)
BVD control measures on service bulls for mixed-age cows (n= 255)	Bulls not used	48	50 (49 ~ 50)	23 (22 ~ 23)	Reference
	No	46	57 (56 ~ 59)	34 (33 ~ 35)	1.72 (0.78 ~ 3.87)
Farmer's awareness of BVD (n= 294)	Yes	161	197 (195 ~ 198)	96 (95 ~ 96)	1.13 (0.59 ~ 2.18)
	Not good	172	178 (177 ~ 179)	89 (88 ~ 90)	Reference
Farmer's awareness of PI animals (n= 297)	Good	122	126 (125 ~ 127)	63 (62 ~ 64)	1.00 (0.63 ~ 1.62)
	No	51	54 (53 ~ 54)	22 (21 ~ 23)	<u>Reference</u>
Farmer's awareness of Trojan cows (n= 289)	Yes	246	251 (250 ~ 251)	130 (129 ~ 131)	<u>1.55 (0.84 ~ 2.92)</u>
	No	203	212 (211 ~ 213)	101 (101 ~ 102)	Reference
	Yes	86	92 (91 ~ 93)	51 (50 ~ 51)	1.33 (0.80 ~ 2.22)

Key: IQR, interquartile range; OR, odds ratio; CI, confidence interval

a: actively infected with BVD.

b: values in parenthesis are interquartile range of number of farms across 1,000 imputed datasets.

Table 3. 2. Multivariable logistic regression of significant farm management risk factors on active infection with BVD in 304 New Zealand cattle farms.

Variable	Level/units	Coefficient (SE)	Odds ratio (95% CI)	P-value
Onto-farm movement of heifers/cows	Never	Reference	Reference	< 0.05 ^a
	Occasionally/rarely	0.03 (0.29)	1.02 (0.58 ~ 1.81)	0.92
	All/most years	0.84 (0.30)	2.21 (1.29 ~ 4.24)	< 0.01
Close animal contacts between farms	Never	Reference	Reference	< 0.05 ^a
	Occasionally/rarely	0.97 (0.36)	2.63 (1.33 ~ 5.41)	< 0.01
	Often	0.33 (0.46)	1.39 (0.56 ~ 3.46)	0.48

Key: CI, confidence interval; SE, standard error.

a: The P-value indicates the significance of given variable.

A total of 286 farmers completed the questions about their awareness of BVD, PI animals, and Trojan cows. Over 40% of them (119/286) answered that their awareness of the disease was good. While most farmers (235/286, 82.2%) reported that they knew what PI animals were, only 85/286 (29.7%) responded that they knew about Trojan cows. Farmers with good awareness of BVD were significantly more likely to know the definition of both PI animals and Trojan cows. When awareness of BVD and PI animals was compared, 117 out of 119 farmers (98.3%) with good awareness of BVD reported knowing about PI animals, while 118 out of 167 farmers (70.7%) with lesser awareness of BVD knew about PI animals (P-value < 0.001). Similarly, 63/119 farmers (52.9%) with greater awareness of BVD understood the concept of Trojan cows, whilst only 22/145 (15.2%) with lesser awareness of BVD did (P-value < 0.001).

In terms of biosecurity, it was common practice for farmers to implement BVD control measures on service bulls both for mating the herd of replacement heifers (134/258; 51.7%) and for mating the herd of mixed-age breeding cows (161/255; 63.1%). Self-reported awareness of BVD was not significantly associated with the likelihood of farmers implementing BVD controls for bulls. Among 253 farmers with complete responses about awareness of BVD and implementing BVD control measures on service bulls for mating the heifers, 56 out of 104 farmers (53.8%) with greater awareness of BVD reported to control BVD on service bulls and 75 out of 149 farmers (50.3%) with lesser awareness of

BVD implemented controls on service bulls for mating the heifers (P-value 0.67). Similar patterns were observed for mating the mixed-age breeding cows, as 71/103 farmers (68.9%) with good awareness of BVD reported that they implemented BVD control measures on bulls and 89/146 farmers (61.0%) with lesser awareness of BVD used BVD-controlled bulls for mating the mixed-age cows (P-value 0.26).

3.4.2. Multivariable regression analysis

The median (and IQR) of farms over 1,000 imputed datasets for each category of variables is summarised in Table 3.1, and the results of the univariable and multivariable logistic regression analyses of risk factors for active BVD virus infections based on the imputed datasets are presented in Tables 3.1 and 3.2, respectively. Although five variables were significant in the univariable analysis (Table 3.1), only the frequency of onto-farm movement of heifers/cows and the frequency of close animal contacts between farms were significantly associated with active BVD virus infection in the multivariable model.

The odds of having active BVD virus infection were approximately 2.2 times greater in farms with onto-farm movement of heifers/cows (95% CI: 1.29 ~ 4.24) compared with farms that had no onto-farm movement of heifers/cows during the last five years, and the odds of having active BVD virus infection were 2.7 times greater in farms with occasional/rare animal contacts between farms (95% CI: 1.33 ~ 5.41) compared with farms that never had close between-farm contacts.

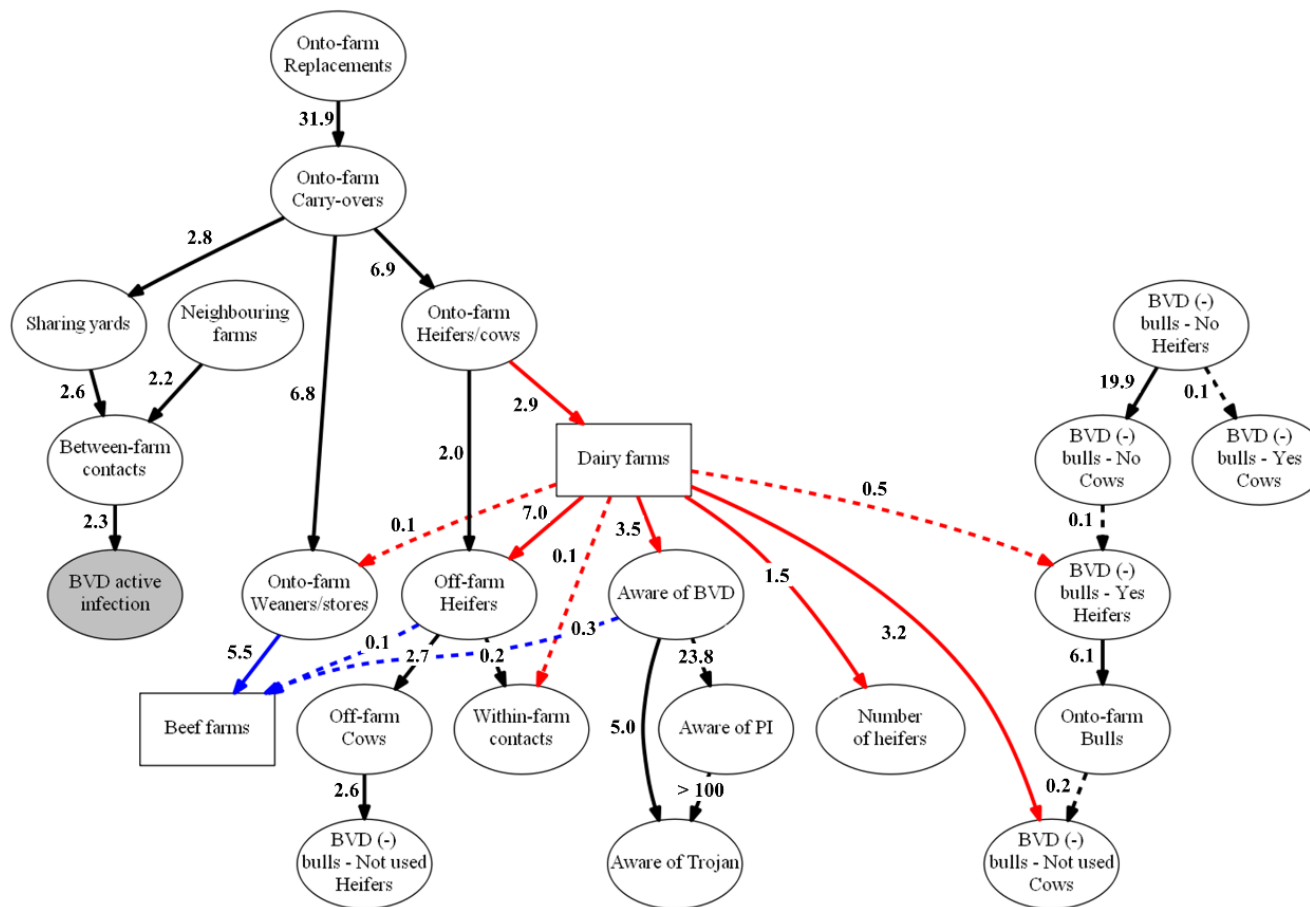


Figure 3. 2. Additive Bayesian network representing significant associations between farm management risk factors on active infection with BVD virus. For each arc, the median odds ratio is reported, and solid and dashed arcs indicate positive and negative associations, respectively. Blue arcs indicate associations with beef farm, and red counterpart describing the one with dairy farms. Active BVD virus infection status is highlighted in grey, and production type (beef and dairy) is described in rectangle node.

3.4.3. Bayesian network analysis

The majority consensus network across 1,000 DAGs had a total of 31 arcs. The arc between the number of replacement heifers and awareness of BVD was dropped from the final network due to a lack of statistical significance. The optimal DAG is described in Figure 3.2 along with the median odds ratio of each arc.

According to the optimal DAG, the odds of being actively infected with BVD virus were significantly higher in herds with more frequent close animal contacts between farms (OR: 2.3, 95% CrI: 1.2 ~ 4.6). Frequency of close animal contacts between farms was also positively correlated with the number of neighbouring farms (OR: 2.2, 95% CrI: 1.1 ~ 4.6) and the frequency of sharing cattle yards (OR: 2.6, 95% CrI: 1.2 ~ 6.1). Other than the frequency of close animal contacts between farms, no variables were directly associated with BVD virus infection status.

Several farm management risk factors were highly associated with production type. Dairy farms had greater odds of onto-farm movement of heifers/cows (OR: 2.9, 95% CrI: 1.8 ~ 4.7) and off-farm movement of replacement heifers (OR: 7.0, 95% CrI: 4.1 ~ 12.6), while beef farms had higher odds of onto-farm movement of weaners/stores (OR: 5.5, 95% CrI: 2.9 ~ 10.7). Close animal contacts within a farm was less frequent in dairy farms (OR: 0.1, 95% CrI: 0.0 ~ 0.4), and awareness of BVD was lower amongst beef farmers (OR: 0.3, 95% CrI: 0.2 ~ 0.5).

With the exception of bulls used for mating, onto-farm movements of different types of cattle were directly or indirectly associated with each other. Farms introducing carry-over cows for grazing tended to introduce other types of cattle as well, including weaners/stores (OR: 6.8, 95% CrI: 3.5 ~ 14.0), heifers/cows (OR: 6.9, 95% CrI: 3.3 ~ 16.5), or replacement heifers (OR: 31.9, 95% CrI: 15.3 ~ 73.6). Management practices applied to the mixed-age breeding cows tended to be the same for the heifers. For example, off-farm movement of the heifers and mixed-age cows was significantly and positively associated (OR: 2.7, 95%

CrI: 1.7 ~ 4.6). Also, implementing BVD control measures on service bulls for mating heifers and mixed-age cows tended to be correlated (OR: 19.8, 95% CrI: 9.2 ~ 45.1).

3.5. Discussion

In the current study, approximately half of unvaccinated beef and dairy farms in New Zealand showed evidence of active BVD virus infection based on the diagnostic criteria of having an S/P ratio greater than 0.75 on the pooled BVD antibody ELISA test. Previous research has shown that farms with a S/P ratio less than 0.75 on pooled serum sample from 15 to 16 animals are extremely unlikely to have PI animals and the probability of having PI animals increased from 14.3% to 58.3% as the S/P ratio increased from 1.00 to 1.50 (Hill, McCoy, et al., 2010). Although it is possible that some of the positive farms were therefore falsely classified as having active BVD virus infection with PI animals, the prevalence of active BVD virus infection on beef farms in our study is in accordance with previous research that reported a prevalence of active infection among beef herds in the North Island of New Zealand was 58.1% (95% CI: 43.4 ~ 72.9%) based on BVD antibody test (W. Cuttance & Cuttance, 2014). Interestingly, the prevalence of active BVD virus infection on dairy farms in our study was higher than previous estimates of 8.6% to 32.0% (varied by region) based on RT-PCR on bulk tank milk (Weir et al., 2016). This could partly be explained by our sampling eligibility criterion, which included only farms that were not implementing any BVD control measure such as vaccination or culling of PI animals. Anecdotal reports from veterinarians suggest that more than 70% of dairy farms are currently implementing at least one control measure for BVD (i.e. vaccination, test and cull, or routine monitoring) compared with less than half of beef farms (Stewart, 2013; Weir et al., 2016). Considering that implementing any BVD control measure is quite common in the dairy industry, the prevalence of dairy farms being actively infected with BVD virus in this study would be overestimated by selection bias. Nevertheless, a key message of this study is that BVD is still endemic in both beef and dairy industries in New

Zealand, and the development of cost-effective BVD control programmes covering the entire country is important to reduce the economic burden of the disease.

As missing values were spread unevenly over the data, the proportion of complete observations in this study was approximately 50% even though overall proportion of data fields with missing data was only 5.8%. Conducting analysis based on complete observation would have not only restricted the use of full extent of the data but also reduced statistical power. It is possible that some of missingness of variables could be correlated, which would invalidate the assumption of “missing (completely) at random” for MICE method. However, given the fact that complete observations comprised just half of the entire dataset, we believe that the benefit of imputing missing values using MICE far outweighed the advantage of using complete cases only even though the assumption may not be fully met.

Cattle farm management in New Zealand is predominantly pasture-grazing with seasonal calving, which may explain why the strength of associations in the multivariable regression model differed slightly from other countries that practice more intensive farming. In systems with intensively housed cattle, close contact over fences could be a relatively minor transmission route and other management factors, including herd size, purchase of animals, or use of artificial insemination, have been suggested to be important risk factors for transmission (Almeida et al., 2013; Solis-Calderon et al., 2005; Van Campen, 2010). In this study, we identified that close animal contacts between farms and onto-farm movement of heifers/cows significantly increased the odds of farms having an active BVD infection. Both between-farm contacts and introducing heifers have been suggested as risk factors for generating PI calves, which is the most important risk factor for BVD transmission. In New Zealand, studies have found that purchasing replacement heifers increased the odds of BVD virus infection in beef herds (W. Cuttance & Cuttance, 2014) and contacts over the fence with neighbouring cattle increased the herd BVD antibody level significantly in dairy herds (Weir et al., 2016). Our findings are in accordance with those previous studies, which

highlights the importance of improving herd biosecurity through measures such as double fencing or movement restrictions to prevent BVD transmission between farms.

Similar to the multivariable regression model, a direct association between active BVD virus infection and the frequency of close animal contacts between farms was identified using the BN. Moreover, a BN uncovered indirect associations between active infection with BVD virus and both the number of neighbouring farms and the frequency of sharing cattle yards via close animal contacts between farms. Importantly, it suggests that the close between-farm animal contacts was actually a mediator variable of other two variables, and that the variable should not be considered for multivariable logistic regression if the other two variables would have been inserted. However, when the number of neighbouring farms and the frequency of sharing cattle yards were considered for the multivariable logistic regression instead of the frequency of close animal contacts between farms, neither of them was significantly associated with active BVD virus infection status (not presented in the study). It indicates that a multivariable regression method would not be the optimal choice to identify indirect risk factors for BVD virus infection in this study, even though it is logical that having multiple neighbouring farms or sharing cattle yards could further increase the risk of BVD transmission by increasing the frequency of potentially infectious between-farm contacts if PI animals were present on other farms (Bitsch et al., 2000; Presi et al., 2011). Therefore, using a BN approach in this study provided a refined and more holistic insight into interrelated risk factors for BVD virus infection in the New Zealand cattle industries by revealing more plausible and statistically significant pathways between them. It may further be implied that local spread could be an important BVD transmission route in pasture-based farms in New Zealand.

Interestingly, we could not find any preventive effect of implementing BVD control measures on service bulls in either analytical approach. This is contrary to findings from a previous New Zealand study, which suggested that vaccinating newly introduced bulls against BVD prevents active infection with BVD virus (W. Cuttance & Cuttance, 2014).

Consequently, using BVD accredited bulls for mating has been a key message to farmers for controlling the disease (Ellison & Weir, 2017), while relatively less emphasis was placed on preventing other ways of BVD virus introduction. According to the optimal DAG, farms with onto-farm movement of bulls tended to implement BVD control measures on service bulls for mating heifers. Therefore, a plausible explanation on the lack of a preventive effect would be that controlling BVD on service bulls has already been a common management feature prior to the study, irrespective of BVD virus infection status. Without controlling for the confounding effects of other routes of BVD transmission, such as allowing between-herd contacts or purchasing replacement heifers with an unknown disease status, the absence of an effect of applying BVD controls on service bulls could be overwhelmed by such factors, and hence confounded to zero.

Another interesting point with regard to the optimal DAG is that there was no direct association between current BVD virus infection status and onto-farm movement of heifers/cows even though this was observed in the multivariable logistic regression. One explanation for this discrepancy between multivariable and multivariate approaches might be a Yule-Simpson paradox that an observed association between two variables can disappear or even be reversed when confounding variable is controlled-for. However, this is a less likely explanation since introducing replacement heifers is a well-known and highly plausible risk factor for BVD and has already been demonstrated in the New Zealand farming context (W. Cuttance & Cuttance, 2014; Weir & Heuer, 2009). Another possible explanation is that the correlation between variables was reduced by generating a consensus network from multiple imputed datasets (Wilson et al., 2013). As the optimal DAG in this study was constructed by retaining only the arcs that were commonly observed over 1,000 structures, the DAG was a conservative network that could have missed some significant associations that did not appear repeatedly across 1,000 networks. It is also possible that the strength of an exposure was reduced by merging two categories into one (e.g. “All/most years” and “Occasional/rarely” became one group) resulting in a weaker association.

Using BN modelling, we identified that beef and dairy farms had different set of risk factors. Dairy farms reported much more frequent onto-farm movement of heifers/cows and off-farm movement of heifers, but fewer within-farm animal contacts between different management groups. This may suggest that dairy farms are at greater risk of introducing BVD virus through cattle movement. In beef farms, onto- and off-farm movement of heifers were less common, but farmers' awareness of BVD epidemiology was also lower, which may make it more difficult to convince farmers of the importance of implementing BVD control measures. Although we could not find any direct or indirect associations between production type and BVD virus infection status, our findings indicate that different risks for BVD virus introduction exist across enterprise types and approaches for farmer education or enhancing biosecurity to control BVD should be tailored to the production type (Gates et al., 2014).

It has previously been shown that the attitude of farmers towards biosecurity can be influenced by their knowledge about the disease (Garforth et al., 2013). However, in our study, we found no significant association between self-reported farmer knowledge of BVD, PI animals or Trojan cows and implementing BVD control measures on service bulls. This may be explained by the fact that New Zealand farmers' awareness of BVD primarily increased after their herds were directly affected by the disease (W. Cuttance & Cuttance, 2014). Furthermore, although many farms in New Zealand have been reported to implement at least one control or preventive measure against BVD, less than 10% of farms implement the full series of interventions (i.e. PI identification and elimination, monitoring BVD status, and BVD vaccination if it is necessary) which is necessary to successfully control the disease (Stewart, 2013). Given that about half of all dairy and beef farms in New Zealand are actively infected with BVD virus, it is prudent to increase efforts to educate farmers about the epidemiology and economic consequences of BVD virus infection to highlight the need for BVD control.

3.6. Conclusions

Based on the results from conveniently sampled cattle farms not using any BVD control measures, close animal contacts between farms appears to be an important BVD transmission pathway in the New Zealand pastoral production system. Herds surrounded by more neighbouring farms or shared cattle yards more frequently are more likely to have between-farm contact. Moreover, onto-farm movement of heifers/cows poses another risk of introducing BVD virus onto the farm. These risk factors can be mitigated with enhanced biosecurity, however, New Zealand farmers may not see the economic benefit of implementing biosecurity measures. Increasing farmer education and awareness of BVD is likely to be a crucial component of future national disease control efforts.

3.7. Acknowledgement

We are grateful to New Zealand cattle farmers and veterinarians who participated in the “Take the BVD Test Challenge” project.

Chapter 4

Estimation of the within-herd transmission rates of bovine viral diarrhoea virus in extensively grazed beef cattle herds

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4.1. Abstract

With the widespread prevalence and significant economic impacts of bovine viral diarrhoea virus (BVDV), many research groups have developed mathematical models to simulate the dynamics of BVDV infections in cattle herds. However, most models use estimates for within-herd BVDV transmission rates that are either based on expert opinion or adapted from other dairy herd simulation models presented in the literature. There is currently little information on the transmission rates for BVDV in extensively grazed beef herds partly due to the logistical challenges in obtaining longitudinal data of individual animal's seroconversion, and it may not be appropriate to apply the same transmission rates from intensive dairy herds given the significant differences in herd demographics and management. To address this knowledge gap, we measured BVDV antibody levels in 15 replacement heifers in each of 75 New Zealand beef breeding farms after their first calving and again at pregnancy scanning or weaning to check for seroconversion during the early pregnancy period. Among these, data from 9 farms were used in a BVDV simulation model to infer the within-herd BVDV transmission rate with an approximate Bayesian computation method. The most probable within-herd BVDV transmission rate was

estimated as 0.11 per persistently infected (PI) animal per day with a 95% highest posterior density interval between 0.03 ~ 0.34. This suggests that BVDV transmission in extensively grazed beef herds is generally slower than in dairy herds where the transmission rate has been estimated at 0.50 per PI animal per day and therefore may not be sufficient to ensure that all susceptible breeding females gain adequate immunity to the virus before the risk period of early pregnancy for generating new PI calves.

4.2. Introduction

Bovine viral diarrhoea virus (BVDV) is recognised for its significant impacts on cattle health, welfare and production worldwide (Anonymous, 2018b). An important epidemiological feature of BVDV is that if susceptible dams are infected in early-mid gestation before the fetus has developed a competent immune system, the fetus will become persistently infected (PI) with BVDV and the resulting calf will shed large quantities of the virus throughout its life-time (Fray et al., 2000). Since PI animals act as the primary reservoir for BVDV transmission in cattle populations, most BVDV control programmes are therefore focused on identifying and eliminating existing PI animals in a timely fashion as well as preventing the creation of new PI animals (Lindberg & Houe, 2005). This can be accomplished through various interventions such as conducting animal- or herd-level diagnostic testing, vaccinating susceptible animals, and improving farm biosecurity (Evans, Pinior, et al., 2019).

To appraise the economic argument for implementing BVDV control measures, several research groups have developed mathematical simulation models to explore BVDV transmission dynamics and its impact on production at varying scales (e.g. in a farm or multiple farms in a region/country) (Cherry et al., 1998; Innocent et al., 1997; Viet et al., 2004). In these models, the infection of a susceptible individual is determined by the force of infection, which is a function of the numbers of PI and transiently infected (TI)

individuals at a given time and the transmission rates (β) for both types of infected cattle. Given the definition of transmission rate is per capita rate at which two individuals have an effective contact (physically close contact with sufficient time that disease transmission could occur if there was an infectious individual) (Vynnycky & White, 2010), PI animals are typically assigned larger β value than TI animals due to their higher viral shedding rates. Consequently, the predictions of simulation models are often highly sensitive to the β value for PI animals (Damman et al., 2015).

An interesting point about BVDV modelling studies with respect to the transmission rates for PI animals is that most studies used parameter values based on either expert opinion or the values assumed by other simulation models (Viet et al., 2007). One of the most commonly used within-herd BVDV transmission rates for PI animals was suggested by Viet et al., (2004) who set the value to 0.5 per PI animal per day (using frequency-dependent assumption) based on other reports (Moerman et al., 1993; Radostits & Littlejohns, 1988). Although it was not estimated from empirical data, the suggested value successfully explained the BVDV spread on a dairy farm, and their method has been reproduced and adopted by subsequent modellers who simulated BVDV spread under their own epidemiological circumstances (Damman et al., 2015; Foddai et al., 2014; Gates et al., 2014; B. J. J. McCormick et al., 2010; Sekiguchi et al., 2018). However, as briefly pointed out in the original paper, a robust transmission rate should be estimated based on longitudinal observations of individual BVDV infection status while considering other management factors, such as production type, population size or density, or herd structure, since those factors can affect BVDV transmission rates within individual herds (Gates et al., 2014; Han, Holter, et al., 2018). Therefore, applying the same β value for simulating BVDV transmission in other populations with substantially different management features such as extensively grazed beef cattle may not be appropriate.

To date, most of the reports on BVDV seroconversion have been based on the dairy industry, whereas only a limited number of studies have focused on the extensively grazed

beef industry due in part to the logistical challenges that make it difficult to collect serial samples of individual animals to measure their seroconversion status (Han, Weir, et al., 2018). The transmission of BVDV in extensively grazed beef herds is expected to be different than in intensive dairy herds since (1) the beef cattle are commonly grazed over a vast area and not often gathered during a year so that the chance of having effective contacts is likely to be lower, while (2) new born calves stay with their dams until weaning (6 ~ 7 months), hence susceptible dams could receive continued exposure to new born PI calves during the breeding period, which likely increases the chance of generating new PI calves in the following calving season.

As part of a larger research programme to address the knowledge gaps around BVDV transmission in extensively grazed New Zealand beef herds, we conducted a panel study to measure BVDV antibody levels in 15 first-calf heifers (replacement heifers immediately after their first calving and before being mixed to the adult cow herd) before and after subsequent breeding period to check for the evidence of seroconversion. Although BVDV transmission rates can be roughly estimated by fitting the change in number of BVDV antibody positive animals over time with a generalised linear regression model (Hage et al., 1996), this method requires detection of all seroconversions based on a series of observations on every animal's BVDV infection status in a herd (Moerman et al., 1993). Given the limited number of sampled animals and sampling occasions in our study, we estimated the transmission rate using an approximate Bayesian computation (ABC) methods. Briefly, ABC is a set of Bayesian methods that infer the posterior distribution of parameters by randomly drawing a sample of the parameters from an initial distribution to simulate data and accepting only the sample of which the simulated data is close enough to the observed data according to a pre-defined distance (or tolerance) (Toni et al., 2009).

Using the seroconversion data from our field studies, the objectives of this current study was to infer the within-herd BVDV transmission rate for PI animals (β_p) in beef breeding

farms using an ABC method while accounting for differences in the herd management structures.

4.3. Materials and methods

4.3.1. Data collection and extraction

A panel study was conducted from September 2017 to September 2018 to estimate the proportion of first-calf heifers that were still susceptible to BVDV prior to their second breeding and to estimate the rates of seroconversion to BVDV during the breeding/early pregnancy period. A total of 75 commercial beef breeding farms from different regions across New Zealand were recruited by convenience sampling through 10 participating veterinary clinics. The selection criteria were that the farms (1) were not currently vaccinating replacement breeding females against BVDV since vaccination is known to interfere with antibody-based screening tests for BVDV, (2) had at least 15 replacement heifers due to calve in the 2017/2018 calving season (generally between August 2017 and November 2017), and (3) were willing to yard animals at two time points for sampling (once prior to the second breeding (i.e. first sampling event) and again at pregnancy scanning or weaning (i.e. second sampling event)).

For each enrolled farm, a participating veterinarian collected blood samples from 15 randomly selected first-calf heifers with calves at-foot before the breeding period and recorded the ear tag IDs for each individual animal. The samples were then transported to a commercial veterinary diagnostic laboratory and analysed individually using a BVDV antibody ELISA test (IDEXX BVDV Total Ab Test; IDEXX Laboratories Inc., Westbrook MA, USA). The samples were considered as BVDV antibody test-positive if the sample to positive (S/P) ratio was > 0.17 which was determined by the laboratory (cut-off value by the manufacturer was 0.2). A survey was administered to farmers at the time of sampling to collect information about general management practices including the date of

management events (breeding, weaning, and calving) and demographic features (herd structure, herd size, and age of heifers when they first calve). A copy of the complete survey is provided in Appendix 2 (see Section 1).

Based on the test results from the first sampling event, BVDV antibody negative heifers were identified, and the participating veterinarians were asked to re-collect blood sample from each negative animal at either pregnancy scanning or weaning. The samples were transported to the same diagnostic laboratory for analysis using the same ELISA test, and the same cut-off value ($S/P > 0.17$) was used to identify whether animals had become BVDV antibody positive. We assumed that a change in status from negative to positive indicated seroconversion to BVDV, so only the results of farms with increased number of BVDV test-positive heifers in the second sampling event were extracted for further study.

4.3.2. Simulation model

In order to estimate the β_P values using an ABC algorithm, we first developed a stochastic individual-based BVDV transmission simulation model to replicate BVDV dynamics in a typical extensively-grazed New Zealand beef herd. For most New Zealand beef farms, weaning occurs at approximately 6 to 7 months of age (Geenty & Morris, 2017). Initial breeding occurs at approximately 14 to 15 months of age with first calving at approximately 24 months of age on most of beef farms, while some farmers breed heifers for the first time at approximately 26 to 27 months old with them calving at approximately 36 months old. The breeding period usually lasts from 8 to 12 weeks. Based on the survey result, most farmers kept the replacement heifer herd separate from the adult breeding cows even after their first calving.

Since the exact mixing date of replacement breeding heifers with adult breeding cows for each eligible farm was unknown, we assumed that the mixing coincided with pregnancy scanning of replacement breeding heifers (and weaning of their first calves) so that the

replacement heifer herd had been grazed separately (except from their calves) for the whole study period since their weaning. Accordingly, the model only included two demographic groups; the replacement heifers and their calves at-foot. We assumed that no other contacts with other groups of animals on the farm occurred during this time period, and BVDV transmission from breeding bulls was not considered given that the farmers reported using BVDV vaccinated or BVDV-free certified breeding bulls.

The simulation model started at the day of weaning for the replacement heifers (day 0) and followed individual animals through their first year of life, initial breeding period, their first calving, the day of first BVDV blood sampling event, second breeding period, and the day of second sampling event (pregnancy scanning or weaning). The total number of heifers and the calendar dates for any management events were taken directly or calibrated from the survey data provided by the farmers to make the models as herd-specific as possible. The calves-at-foot group was only modelled from birth until weaning.

BVDV status of animals in the simulation model was either; immune via maternal antibody (M; only for the calves at-foot), susceptible (S), transiently infected (T), persistently infected (P), and recovered (R). BVDV transmission was assumed to be frequency-dependent under homogeneous mixing. Since we did not assess the BVDV status of the entire herd nor conduct any antigen testing to confirm the presence of PI animals, we could not actually determine when the tested heifers were exposed to BVDV or the source of their exposure. Given that BVDV is endemic in New Zealand (Han, Weir, et al., 2018), we therefore assumed that a random proportion of replacement heifers in each farm (μ_i) were recovered and seropositive to BVDV at the time of weaning (day 0) due to the anamnestic response (if calves with enough maternal antibody were exposed to BVDV, they show a higher immune response to BVDV when they are re-exposed to the virus after the depletion of maternal antibody) (Platt et al., 2009; Ridpath et al., 2003). We then simulated from the day of weaning (day 0) and introduced BVDV via PI animals at a random point of time between day 0 and the day of first sampling event (τ_i) for each farm.

The introduction of PI animals in the simulation model was conducted by converting a random proportion of heifers as PI (ρ_i). The introduction of BVDV via TI animals was not considered as they have been demonstrated to have a limited role in BVDV transmission (Niskanen et al., 2000). Also, it was assumed that the introduction of PI animals occurred at a single time only, and that 50% of the introduced PI animals had been removed annually from the point of introduction (Innocent et al., 1997).

The probability of BVDV infection for each susceptible animal (*Prob*) at a given day was;

$$Prob = 1 - e^{-\lambda_t}$$

$$\lambda_t = \beta_P \frac{P_t}{N} + \beta_T \frac{T_t}{N}$$

where λ_t is the force of BVDV infection at day t , P_t and T_t are the number of PI and TI animals, respectively, at day t , N is the herd size, and β_T is the within-herd BVDV transmission rate for TI animals. Since our aim was to estimate β_P , β_T was assumed to be proportionally (0.05) lower than β_P with the proportion being calibrated based on other studies (Smith et al., 2014; Viet et al., 2007). One should note that the meaning of transmission rates above is actually the number of effective contacts per infectious individual per day (i.e. effective contact rate), which is different from the classic definition of transmission rate (i.e. per capita rate at which two individuals have an effective contact per unit time) (Vynnycky & White, 2010). However, we used the effective contact rate and transmission rate interchangeably for the coherence with other BVDV modelling works. TI animals were assumed to recover and became seropositive after x days of infection, where x was randomly chosen between 10 and 20 days (Liebler-Tenorio et al., 2004; Müller-Doblies et al., 2004; T. Sandvik et al., 1997). Once recovered, animals were assumed to have protective immunity against BVDV for the remainder of the simulation given the short duration of simulation.

Each simulation started by establishing a female herd of N_i weaned animals, where N was the reported herd size of replacement breeding heifers via survey for farm i . Puberty in heifers was assumed to start at 380 days of age (McNaughton et al., 2002) with oestrus cycles occurring every 18 to 24 days thereafter until the animal became pregnant. When a pregnant heifer calved or had an abortion, the next oestrus was assumed to occur after 15 to 49 days (Olds & Seath, 1953) and assumed to be a silent oestrus. When the simulation reached to the start date of initial breeding period which was calibrated based on the date of breeding and age of heifers to calve in the survey, heifers were mated with bulls and only the heifers in oestrus were conceived with the probability of 62.0% during the breeding period (McFadden et al., 2005). The gestation period was 281 days and the natural abortion rate was set at 0.0001 per day over the whole gestation period to achieve a total abortion rate of 3.5% for the season (Hickson et al., 2012; Weston et al., 2012). For each new born calves from heifers, 0.0009 per day of natural mortality rate was applied until weaning (86.7% of calves surviving to 180 days) (C. A. Morris et al., 1986).

The probability of abortion of replacement heifers due to BVDV infection was varied by the days of pregnancy at the time of infection (Table 4.1). Likewise, the BVDV infection status of new born calves was determined by the stage of pregnancy the heifers were in when they became infected with BVDV. The period of maternal BVDV antibody protection for individual calves born to recovered replacement heifers was randomly selected from a normal distribution with the mean and standard deviation of 155 and 31 days, respectively (Coria & McClurkin, 1978; Kendrick & Franti, 1974) (see Section 2 of Appendix 2).

At the day of first sampling event, 15 individuals were selected randomly from the simulated heifers and BVDV ELISA antibody test-positivity of each individual was probabilistically determined based on each animal's BVDV infection status, test sensitivity (0.99), and specificity (0.81) (Beaudeau et al., 2001). Each simulation ended when it reached the day of second sampling event, where similar to the field studies, any of the

individual animals that tested negative for BVDV antibodies were re-sampled to determine whether they had seroconverted. An illustration of the simulation is provided in Figure 4.1 and the definition and values for the model parameters are shown in Table 4.1.

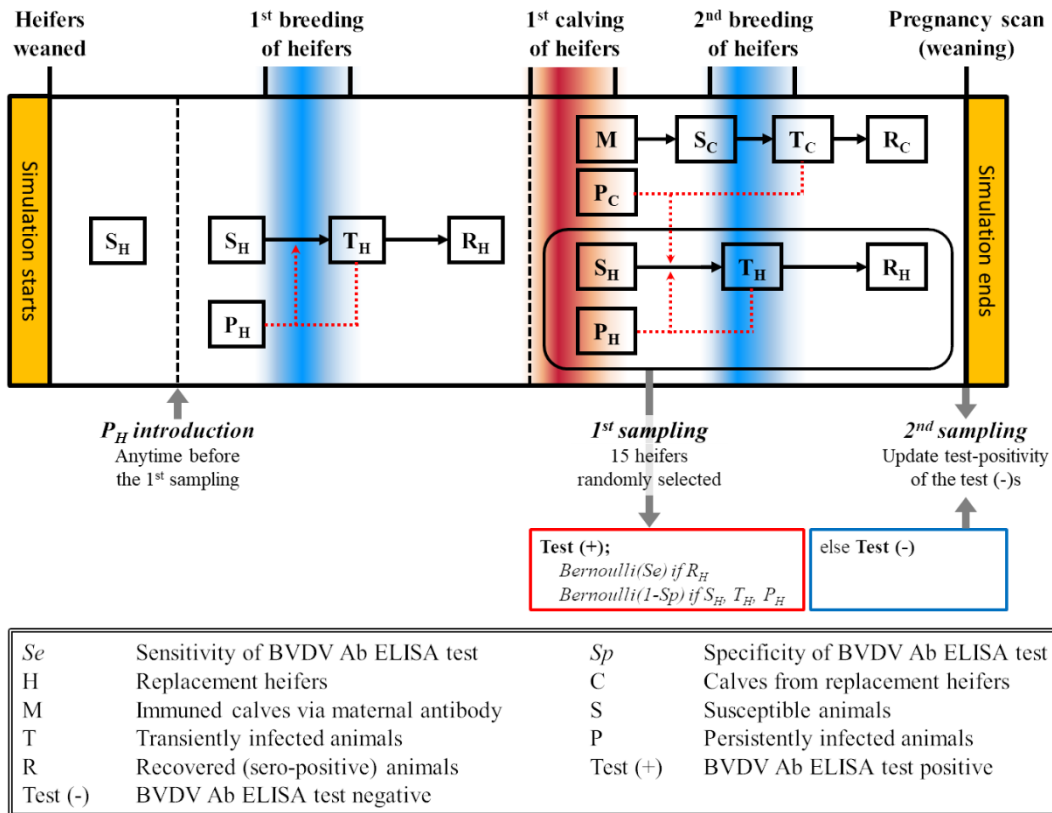


Figure 4. 1. Description of within-herd simulation for estimating within-herd BVDV transmission rate. Blue and red shaded areas indicate the breeding and calving periods, respectively. Red dotted arrow represents the contribution to the force of BVDV infection.

Table 4. 1. Definition and value of parameters used for within-herd BVDV transmission model.

Parameter definitions	Value	Reference
Oestrus cycle (day)	$U(18, 24)$	McNaughton et al., (2002)
Probability of conception	0.62	McFadden et al., (2005)
Natural abortion rate	0.0001 per day	Hickson et al., (2012; Weston et al., (2012)
Natural calf mortality rate until weaning	0.0009 per day	C. A. Morris et al., (1986)
β_p : Within-herd BVDV transmission rate (per PI animal per day)	To be estimated	
β_T : Within-herd BVDV transmission rate (per TI animal per day)	$\beta_p \times 0.05$	Smith et al., (2014); Viet et al., (2007)
Infectious period of TI animals (day)	$U(10, 20)$	Liebler-Tenorio et al., (2004; Müller-Doblies et al., (2004); Sandvik et al., (1997)
Period of maternal BVDV antibody protection (day)	$N(155, 31^2)$	Coria and McClurkin, (1978); Kendrick and Franti, (1974)
Mortality rate of PI animals	0.0019 per day	Damman et al., (2015); Innocent et al., (1997)
Probability of abortion if infected during early pregnancy (day 0 ~ 41)	0.8	Ezanno et al., (2007); Moennig and Liess, (1995)
Probability of abortion if infected during mid-pregnancy (day 42 ~ 150)	0.25	Ezanno et al., (2007); Moennig and Liess, (1995)
Probability of calving PI animal if infected during mid-pregnancy	0.934	Ezanno et al., (2007); Fray et al., (2000)
Probability of calving an immuned calf if infected during mid-pregnancy	0.033	Ezanno et al., (2007); Fray et al., (2000)

Key: $N(\cdot, \cdot)$, normal distribution(mean, variance); $U(\cdot, \cdot)$, uniform distribution(lower limit, upper limit).

Table 4. 2. Prior distribution of parameters estimated using ABC-SMC algorithm.

Parameter	Prior distribution
Within-herd BVDV transmission rate for PI animals (β_p)	Beta (1.18, 1.54)
Initial proportion of BVDV seropositive animals (μ) *	Uniform (0, 1)
Proportion of introduced PI animals (ρ) *	Beta (1.14, 7.79)
Day of PI animals being introduced (τ) *	Uniform (0, k)

* Sampled values varied by herd.

Key: k, number of days from weaning of replacement heifers to the day of first sampling event.

4.3.3. Estimation of within-herd BVDV transmission rates

Along with β_p , there was no information about the initial proportion of BVDV seropositive animals (μ_i), the proportion of introduced PI animals (ρ_i), and the day of PI animals being introduced (τ_i) in each study herd i . Given the partial observation of whole BVDV transmission in this study with intractable likelihood of the simulation model, we estimated parameters using an ABC method. Specifically, we chose an approximate Bayesian computation-sequential Monte Carlo (ABC-SMC) algorithm to achieve the computational efficiency (Beaumont, 2019; Toni et al., 2009). In the ABC-SMC, parameters are inferred through a series of estimation sequences (or SMC sequence), each of which has a decreased distance threshold (ϵ) than its previous sequence. A SMC sequence starts with selecting a parameter value based on the perturbation of estimated values in the previous sequence (or the random sampling of a prior distribution in case of the first SMC sequence), and applying the selected value to a model to generate data (i.e. simulated data). The difference between observed and simulated data is measured by comparing the summary statistics of each data (i.e. distance), and only the value showing the distance less than a threshold of the sequence is retained. Current SMC sequence ends when a number of values (i.e. particle) are collected, and the whole process repeats until the particles of final SMC sequence are collected. The distribution drawn by the particles of final sequence is a marginal posterior distribution, which approximates the posterior distribution of the parameter given the observed data. Since it is a modeller's choice which

summary statistics to use to measure the distance, the chosen summary statistics should contain sufficient information about the parameter because it affects the posterior distribution of the parameter with only the reduced information of each data (Bertorelle et al., 2010). The detailed algorithm for the ABC-SMC is provided in Appendix 2 (see Section 3).

We assumed that all study herds shared the common distribution of β_P while other parameters varied by the herds. For the prior distribution of within-herd BVDV transmission rate for PI animals ($\pi(\beta_P)$), we assumed that the value was less than 0.5 with mode of 0.25 (with 60% certainty). The mode of the prior distribution was set to achieve 60% of susceptible animals being infected by a PI animal in a year (Gunn et al., 2004), however, the distribution was still wide enough to cover various values between 0 and 1. A uniform prior between 0 and 1 was used for the initial proportion of BVDV seropositive animals ($\pi(\mu_i)$). The prior distribution of the proportion of introduced PI animals ($\pi(\rho_i)$) followed a beta distribution, with a certainty of 50% that the chosen value was less than 0.1 (with the mode at 0.02). The prior of the day of PI animals being introduced ($\pi(\tau_i)$) followed a uniform distribution, as it was randomly selected between 0 (i.e. PI animals had been co-grazed with the heifers from their weaning onward) and the number of days from day 0 to the day of first sampling event (i.e. PI animals were introduced at the day of first sampling event). All priors were assumed to be independent.

For the perturbation kernel, a component-wise Gaussian kernel with the variance as 0.68 times of the variance of particles in the previous SMC sequence was used (Ellen Brooks-Pollock et al., 2014). We used two summary statistics to estimate the distance (D_k) between the observed and simulated data; (1) the observed (Obs_1) and simulated (Sim_1) number of test positive heifers in the first sampling round (D_1), and (2) the observed (Obs_2) and simulated (Sim_2) number of test positive heifers (among those were test negative at the first round) in the second sampling round (D_2). Each summary statistics was calculated as below;

$$D_k = \sqrt{\sum_{i=1}^n (Obs_k^i - Sim_k^i)^2}$$

where n was the number of herds in this study. For each SMC sequence, the simulation model was iterated until 2,000 acceptable particles were generated, and the acceptable distance thresholds were set as 50th percentile of the distances of accepted particles in the previous sequence. The marginal posterior distribution of each parameter was achieved after 15 SMC sequences.

4.3.4. Validation of the estimated parameters

Once the posterior distributions were retrieved, we inspected the fitness of the estimated parameters by reproducing BVDV transmission within study herds with the estimated parameter values while considering management features of each herd. To do this, we randomly sampled 2,000 sets of parameter values from the respective posterior distributions, and applied to the simulation model to draw a distribution of simulated summary statistics (Sim_1 and Sim_2). We then visually examined whether the observed statistics located within a 95% prediction interval of the simulated summary statistics (Bertorelle et al., 2010; Sunnåker et al., 2013).

We also conducted a sensitivity analysis to investigate the impact of using different prior distribution of unknown parameters. Before the sensitivity analysis, we preliminarily examined which parameter affected the most to the summary statistics in an individual farm. It showed that the statistics were the most sensitive to the values of μ and ρ , while other parameters had only a limited impact (see Section 4 of Appendix 2). Since there was a lack of information to estimate an informative $\pi(\mu_i)$ in this study, we tested two different distributions (e.g. uniform distribution between 0 and 1, and normal distribution with mean of 0.3 and standard deviation of 0.1) of only the $\pi(\rho_i)$ for the sensitivity analysis. The

disease simulations were implemented in the C programming language and ABC-SMC algorithm was run in R (R Core Team, 2018).

4.4. Results

4.4.1. Longitudinal data

Overall, blood samples from 1,116 individual heifers on 75 farms were collected at the first sampling event, and, of these, all heifers from 12 farms were BVDV antibody positive and 729 heifers from 63 farms tested negative for BVDV antibodies. Farms where all sampled heifers tested positive were excluded from follow-up, and we were only able to obtain the second blood samples from 673/729 heifers (92.3%) located on 55/63 farms due to logistical issues with being able to yard the same animals for re-sampling (one to five heifers were unavailable to follow-up on 19 of the 63 farms).

From the 55 re-sampled farms, all heifers from 25 farms were BVDV antibody test-negative for both sampling events, indicating no circulation of BVDV before and during the study period. Test results from 17 farms showed no additional BVDV antibody test-positive heifers possibly due to early culling/moving of PI animals, and only the 13 remaining farms indicated BVDV seroconversion. The data from these 13 farms were retained for further use in this study, however, records from 4 of these farms were discarded either because some of the animals were later found to have been BVDV vaccinated or because the second sampling event took place outside of the time-window specified for the study. Overall, test results of 133 heifers from 9 farms were used to estimate β_p values. Descriptive statistics on the number of sampled and test-positive heifers at each sampling, and other management features, such as herd size, breeding period, and day of both sampling events, for each farm are provided in an additional file (see Section 5 of Appendix 2).

4.4.2. Parameters estimation

The posterior distribution of β_P is illustrated in Figure 4.2 and the evolution of β_P values over 15 ABC-SMC sequences is provided in Appendix 2 (see Section 6). The mode of posterior β_P was 0.11 with a 95% highest posterior density (HPD) interval between 0.03 and 0.34 per PI animal per day. The posterior distributions of estimated μ , ρ , and τ are illustrated in Table 4.3 and Figure 4.3, and the evolution of parameter values along the SMC sequences is provided in Appendix 2 (see Section 6). The distributions of μ and ρ varied between herds, with the mode of μ ranged between 0.07 and 0.84, and the mode of ρ between 0.03 and 0.16. The estimated day when PI animals were first introduced varied by herd as well. The most probable (mode) day for Farms 1, 3, 6, and 7 occurred during or immediately after the first calving period, while the day for Farm 2 and 9 were during the first breeding period. The estimated day for the remaining farms were widely dispersed between the day of heifers being weaned and the first breeding period.

4.4.3. Validation

The observed and simulated summary statistics for each study herd are provided in Appendix 2 (see Section 6). For both Sim_1 and Sim_2 , observed summary statistics for all 9 herds were located within a 95% prediction interval of the simulated summary statistics. Most of the observed summary statistics were located near the mode of simulated summary statistics, however, the observed Sim_2 of some but not all herds were located near the margin of 95% prediction interval (see Section 6 of Appendix 2). In the sensitivity analysis, some but not all parameters showed different posterior distributions when either a uniform or normal distribution of $\pi(\rho_i)$ were used. Especially, the posterior distribution of ρ_i was generally affected by the prior distribution of itself, indicating that the parameter ρ_i in our model was not identifiable in general (see Section 7 of Appendix 2).

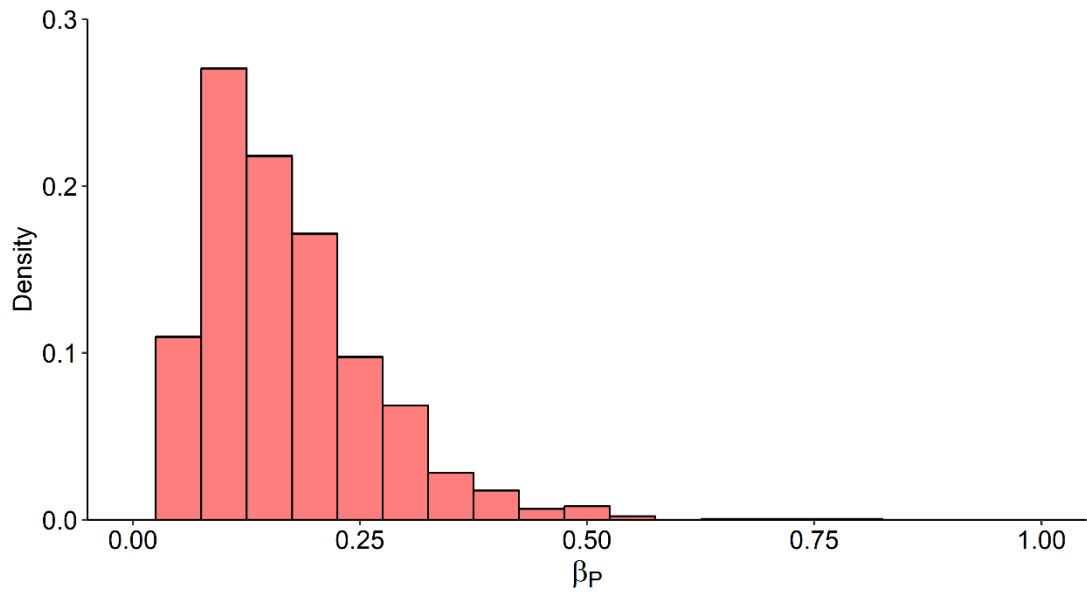


Figure 4. 2. The posterior distribution of within-herd BVDV transmission rates for PI animals (β_P). The mode was 0.11 with a 95% highest posterior density interval between 0.03 and 0.34.

Table 4. 3. Mode (95% highest posterior density interval) of the estimated initial proportion of BVDV seropositive animals (μ), the proportion of introduced PI animals (ρ), and the day of PI animals being introduced (τ) for each studied farm.

Farm	μ	ρ	τ
1	0.32 (0.00, 0.58)	0.11 (0.01, 0.32)	546 (359, 590)
2	0.53 (0.05, 0.93)	0.10 (0.00, 0.42)	252 (13, 499)
3	0.07 (0.00, 0.29)	0.03 (0.00, 0.40)	525 (380, 533)
4	0.62 (0.10, 1.00)	0.15 (0.00, 0.39)	210 (24, 530)
5	0.80 (0.07, 0.98)	0.14 (0.00, 0.39)	378 (18, 866)
6	0.26 (0.00, 0.56)	0.04 (0.00, 0.41)	588 (225, 625)
7	0.08 (0.00, 0.25)	0.16 (0.03, 0.37)	945 (913, 950)
8	0.48 (0.00, 0.88)	0.12 (0.00, 0.40)	357 (53, 597)
9	0.84 (0.09, 0.98)	0.15 (0.01, 0.38)	294 (9, 543)

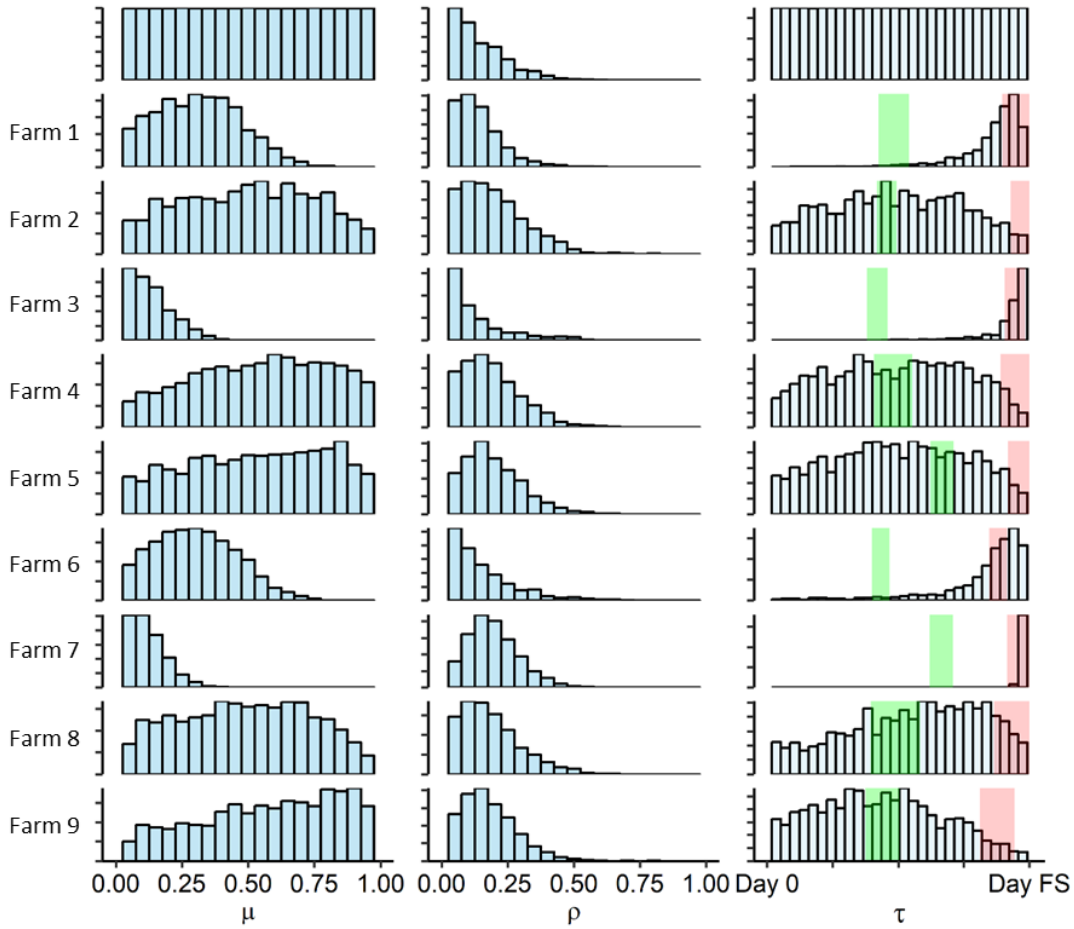


Figure 4. 3. The posterior distribution of the estimated parameters for each herd. Left: the initial proportion of BVDV seropositive animals (μ), Middle: the proportion of introduced PI animals (ρ), Right: the day of PI animals being introduced (τ). The distributions of the first row illustrate the prior distribution of each parameter. Green and red shaded areas indicate the first breeding and calving period, respectively. Day FS indicates the day of first sampling.

4.5. Discussion

Using an ABC-SMC algorithm, we estimated the within-herd BVDV transmission rates for PI animals based on field data collected from extensively grazed beef herds. Although the ABC method has been used to estimate the transmission rate of other animal diseases, such as bovine tuberculosis (Ellen Brooks-Pollock et al., 2014), Peste des petits ruminants (Fournié et al., 2018), and African swine fever (Guinat et al., 2018), our study is the first attempt to use this method for estimating β_P values to the best of our knowledge. This is most likely due to the paucity of BVDV seroconversion data, especially in extensively

grazed beef herds. Other studies have used a commonly reported β_P value from intensively managed dairy herds to explain BVDV spread (Damman et al., 2015; Foddai et al., 2014; Gates et al., 2014; B. J. J. McCormick et al., 2010; Sekiguchi et al., 2018). However, more efforts should be made to estimate robust β_P values based on empirical data that take into consideration other herd-level management factors that can affect BVDV transmission.

Estimation of β_P values was based on only 9 out of 75 initially recruited farms, and we acknowledge the possibility of selection bias in this study. While it is possible that the 12 farms that were not followed-up at the second sampling event had higher β_P values than the estimated values in this study, it is also plausible that the heifers on those farms had been exposed to the virus for much longer time periods (e.g. co-grazed with PI calves from birth) or that there were a larger number of PI animals compared to other farms that could have led to more efficient within-herd BVDV transmission. The fact that none of the heifers on 17 of the re-sampled farms seroconverted to BVDV was most likely due to the death or removal of PI animals before the first sampling event rather than very low BVDV transmission rates for co-mingled PI animals. Even under extensive beef farming conditions in New Zealand, we believe that the duration between sampling events (127 to 177 days; see Additional file 5) was sufficient for susceptible heifers to have effective contacts with PI animals if PI animals were actually present in the herd. In future studies with more available resources, it would be useful to track the serological status of replacement heifers from an earlier age and perform whole herd BVDV testing to determine the presence of PI animals as this would allow more accurate estimation of β_P .

Different values of β_P have been suggested in the literature, but cannot be directly compared given that each group used different modelling structures and assumptions. However, the values can be indirectly compared by estimating the cumulative numbers of infected animals after one year of a PI animal being introduced to 100 susceptible animals in a closed herd while ignoring transmission by TI animals (Viet et al., 2007). Using the previously reported β_P values, the cumulative numbers of infected animals ranged from

approximately 30 (Gunn et al., 2004; Innocent et al., 1997) to 100 (Cherry et al., 1998; Pasman et al., 1994; Sørensen et al., 1995) with some intermediate estimations of 60 (Gunn et al., 2004) and 84 (Viet et al., 2004). When applying the same setting to β_P values estimated in this study, it was equivalent to seroconvert 33 susceptible animals (mode) with a range between 10 and 71 animals (95% HPD interval). Our estimates are somewhat lower than those based on the previously reported β_P values, indicating that using previous β_P values to extensively grazed New Zealand beef farms may overestimate the transmission of BVDV. Given the fact that previous β_P values were estimated based on either expert opinion or data from dairy herds, our estimate of β_P would be more suitable for modelling BVDV in pasture-based beef industries.

In addition to estimating β_P values, we also discovered several interesting features of BVDV epidemiology from this study. First, the most probable proportion of introduced PI animals was estimated between 3% and 16%, which varied between farms. An anecdotal report showed that the prevalence of PI animals within a cattle herd in New Zealand varied between 2% and 12% (Thompson, 2005), and so our estimate of proportion of introduced PI animals was in accordance with the previously reported range. Second, by considering the management feature, such as breeding and calving events, of replacement heifer herds, we were able to infer BVDV transmissions from introduced PI animals as well as any new born PI calves. In particular, Farms 1, 3, 6, and 7 showed that the estimated day of PI animals being introduced was during or immediately after the first calving period. This indicates that the introduced PI animals were actually new born PI calves either from heifers infected during early-mid pregnancy or potentially other purchased pregnant replacement heifers carrying a PI calf (Trojan dam). On the other hand, the most probable day of PI animals being introduced for Farms 2 and 9 was during the first breeding period, possibly indicating young PI replacement heifers were purchased for breeding. Unfortunately, we did not have detailed information on each farm's purchasing history to confirm our speculations. However, these results still highlight that the ABC-SMC

algorithm is a useful tool to infer unknown BVDV-related parameters, such as the timing and type of BVDV introductions, even with a limited amount of data.

In our simulation model, β_T was assumed to be the product of β_P and a proportion (0.05) that quantifies the reduced chance of the effective contacts with TI animals compared to PI animals, given that TI animals shed lower quantities of virus and likely require longer periods of time and closer contacts to cause transmission. Although we relied on other studies to calibrate this proportion, our preliminary investigation indicated that the proportion would have no marked impact on the estimated β_P values, further implying that the uncertainty about β_T could be negligible for estimating β_P . We also conducted a sensitivity analysis using a uniform and normal prior distribution of ρ_i . The sensitivity analysis showed that the posterior distributions of some parameters were highly affected by the prior distribution of ρ_i (see Section 7 of Appendix 2). It indicates that some parameters, especially ρ_i , were not identifiable under the current model setting, possibly due to the lack of data and/or complex model structure, and the accurate estimation of parameter values in this study was generally dependent upon the precise information of $\pi(\rho_i)$. Although we argue that our estimated parameter values were valid considering that the prior of ρ_i in this study well-matched to the reported prevalence of new-born PI calves under New Zealand farming context (Thompson, 2005), more observations — the number of tests as well as the sample size of each test — per farm should be warranted in the future study to mitigate the non-identifiability issue.

We consider several potential limitations to the study methods that could influence the accuracy of our estimates for β_P . First, the ABC-SMC algorithm was applied to data on only two sequential BVDV antibody ELISA tests for the subset of sampled animals in each herd. Even though we only accepted the parameters which were able to simulate data that closely matched the observed test results while adjusting for the test sensitivity and specificity, two sampling events may be regarded as insufficient. A lack of enough data points as well as errors introduced by the random sampling of heifers in this study might

have increased the variability of the posterior distribution of not only β_P but also other parameters (Gelman et al., 2013). More precise estimation could be achieved by conducting future longitudinal studies that sample more animals at more frequent intervals and test individual animals to confirm the presence and identity of PI animals. However, the logistical challenges of conducting sampling on extensively grazed beef herds will still remain as an obstacle, particularly since beef farmers are often reluctant to yard dams with calves at-foot due to the risk of injury to calves (Han, Weir, et al., 2018). Second, in our simulation model, we ignored any virus introduction from different management groups within the same farm or from other external sources (e.g. co-grazed sheep), which may have led to an overestimation of β_P values. However, considering that tested heifers were grazed extensively, it is less likely that the heifers had frequent effective contacts with animals from different herds. Also, our current research about BVDV transmission between co-grazed cattle and sheep in New Zealand beef farms suggests that the transmission between two species is a rare event (unpublished). Finally, we assumed that PI animals were introduced at only one time point. However, based on unpublished observations from the national animal movement data in New Zealand and data from the management surveys completed by farmers, it is unlikely that many beef farmers are purchasing replacement heifers at multiple time points. The most likely alternate source of PI animals is through the birth of PI calves from susceptible animals that were exposed to BVDV during pregnancy. Therefore, we believe that the posterior distribution of β_P in this study provides valid estimates of the most likely values.

Overall, the study findings have important clinical implications for the control of BVDV especially since the estimated β_P values indicated that BVDV transmission is likely to be slow in extensively grazed beef herds. A similar finding was reported in a New Zealand cattle farm by Thompson, (2005), who observed that approximately 20% of animals remained susceptible to BVDV even after being directly co-grazed with PI animals for up to 600 days. These findings strongly suggest that using PI animals as a “*natural*” BVDV

vaccination source (i.e. purposely exposing PI animals to susceptible animals to replicate the protective effect of BVDV vaccination) is likely to fail in the beef herds. On the contrary, this practice has a high risk of generating new PI calves in the next calving season with the risk of causing a BVDV outbreak in such extensively grazed beef cattle farms. Under these circumstances, effective prevention of BVDV should be based on test-and-cull of identified PI animals, vaccination of susceptible animals, and improvement of farm biosecurity (Evans, Pinior, et al., 2019).

4.6. Conclusions

By measuring the BVDV seroconversion of first-calf heifers during the mating period, we inferred that the within-herd BVDV transmission rate of a PI beef animal in extensively grazed herds was 0.11 per day which is lower than the one in dairy herds or herds with intensive farming system. This finding indicates that the calving of PI animals in extensively grazed New Zealand beef farms would not confer adequate herd immunity for naïve dams that co-grazed with new-born calves. This slow dynamics of BVDV transmission should be considered in establishing BVDV control strategies for the New Zealand cattle industries.

4.7. Acknowledgements

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Chapter 5

Economics of bovine viral diarrhoea virus control in pastoral dairy and beef cattle herds

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5.1. Abstract

Bovine viral diarrhoea virus (BVDV) is a prevalent pathogen in the New Zealand cattle industries, yet few studies to date have evaluated the economics of BVDV in pastoral dairy and beef herds to help inform management decisions. To address this knowledge gap, we developed stochastic individual-based simulation models to represent the transmission dynamics of BVDV in typical spring-calving dairy and beef farms in New Zealand. The models conservatively estimated the direct losses due to a BVDV outbreak at NZ\$ 22.22 and NZ\$ 41.19 per mixed-age cow per year for a naïve dairy and beef farm, respectively, over a 5-year period. The greatest economic impacts for the dairy farm occurred when a persistently infected replacement heifers joined the lactating cow group and caused transient infection of cows to drop in milk production, whereas the greatest impacts for the beef farm was through the loss of fattening stock for sale due to lowered pregnancy rates. Various combinations of diagnostic testing, vaccination, and biosecurity measures were then explored to evaluate the cost-efficiency of different strategies for controlling BVDV at the farm-level. We identified that the annual testing of breeding calves with or without double fencing on dairy farms and the vaccination of breeding animals with annual booster

on beef farms were economically beneficial. Providing farmers with the estimates of economic impacts of BVDV in their herds may further encourage the uptake of control measures, but close collaboration with a veterinarian to determine the optimal strategy for their unique farm circumstances is still required.

5.2. Introduction

Bovine viral diarrhoea virus (BVDV) causes a wide spectrum of subclinical and clinical manifestations in affected cattle including infertility, abortion, reduced milk production, diarrhoea, and immunosuppression (Houe, 1999). In New Zealand, BVDV has been endemic since at least the 1960s (Salisbury et al., 1961), and, according to recent studies, nearly half of New Zealand cattle farms have evidence of recent exposure to BVDV with an estimated 10% of farms possibly harbouring persistently infected (PI) animals (Gates, Han, et al., 2019; Han, Holter, et al., 2018). Due to its high prevalence in New Zealand, BVDV is believed to cause significant economic losses for the cattle industries as a whole, although less is known about the impacts of BVDV in individual herds, particularly with extensive beef farms where it is more logistically challenging to yard cattle at the appropriate times to collect data (Han, Weir, et al., 2018).

Currently, there is no compulsory national BVDV control programme in New Zealand, and any uptake of control measures such as testing and vaccinating replacement breeding cattle is at the discretion of individual farmers (Han, Weir, et al., 2018). It has been reported that fewer than 10% of New Zealand dairy farmers routinely test for and cull identified PI animals (Stewart, 2013), although as many as 65% of dairy farms may be performing routine annual bulk tank milk testing (Gates, Han, et al., 2019). In the beef sector, less than 30% of New Zealand farmers had tested their herds for BVDV during the last 5 years, which is believed to be due to the low perceived impact of BVDV or the high perceived cost of BVDV control (Gates, Evans, et al., 2019). Internationally, it has been demonstrated

that farmer's implementation of BVDV control depends on a variety of factors with economic incentives being one of those motivations (Heffernan et al., 2016). Therefore, providing New Zealand farmers with information about the direct losses attributable to BVDV and the cost-benefit of BVDV control could help influence their biosecurity decisions.

Given the time and cost of conducting field studies to evaluate the impact of BVDV and the difficulties of accounting for the effects of different control measures in cattle populations on different scales (i.e. farm-level, regional, or national), simulation modelling approaches are often used as an alternative method (Damman et al., 2015; Sekiguchi et al., 2018; Smith et al., 2014). While there have been numerous published BVDV modelling studies in the literature, the reported results such as the production loss or efficiency of control often cannot be directly applied to other cattle industries due to the difference in herd sizes, industry demographic structure, or efficacy of intervention measures under different farming system constraints (Damman et al., 2015; Smith et al., 2014). To provide more accurate information to farmers, BVDV simulation models should reflect the unique farming characteristics of the target population (e.g. seasonal calving vs. year-round calving, or intensive vs. extensive farming).

To the best of our knowledge, there have been no published simulation studies to date evaluating the epidemiology and economics of BVDV control in New Zealand pastoral dairy and beef herds. Using novel stochastic individual animal-based BVDV simulation models, the aims of this study were to (1) estimate the direct losses due to a BVDV outbreak on New Zealand dairy and beef farms, and (2) investigate the cost-benefit of BVDV control strategies for farms of each production type.

5.3. Material and methods

We first developed two stochastic individual-based BVDV transmission simulation models to replicate the dynamics of BVDV in typical spring-calving New Zealand dairy and beef farms. The models contained (1) a demographic component describing the management group and events on farm, (2) a disease component describing the spread of disease between cattle on the farm, (3) a production component describing the economic outputs of the farm, and (4) a control component evaluating the cost-benefit of different BVDV control strategies. The models operated in discrete time steps with units of one day. For simplicity, we assumed that farms aimed to maintain a constant number of mixed-aged cows in the main breeding herd (T_M) with target numbers of 400 and 150 for dairy and beef farms, respectively, based on the representative herd size reported for each industry (Gates, Evans, et al., 2019). Details used in the estimation of the model parameter values are provided in Section 1 of Appendix 3.

5.3.1. Demographics of dairy farms

The planned start day of calving (PSC) for most New Zealand dairy farms is between mid-July and early August with the calving season typically lasting for approximately 9 to 12 weeks. New born calves are separated from their dams within their first 24 hours, and a proportion of female calves are retained as replacement heifers, particularly those that were born earlier in the calving season and conceived by artificial insemination. The rest of the female calves and virtually all male calves are sold to either abattoirs (i.e. bobby calves) or beef farms after 4 days old. The planned start day of mating (PSM) occurs approximately 84 days after the PSC to keep the herd on a consistent annual calving cycle (Anonymous, 2018a). Another unique characteristic of the New Zealand dairy farming system is that replacement heifers are commonly moved to a distant location for grazing and breeding (i.e. off-site grazing), and returned before their first calving to be joined to the mixed-aged cow herd.

In our model, the cattle population for a dairy farm was divided into 4 management groups: calves (C: from birth until 4 days old), young replacement heifers (YH: from 4 days old to moving off-site for grazing), breeding replacement heifers (BH: heifers at the off-site grazing location until returning and joining the mixed-age cow herd), and mixed-age cows (MA: all breeding female cows that have delivered at least one calf). In this study, the PSC was August 1st and all calves were separated from their dams when they were born and joined C group. All male calves were culled at 4 days old, and the first 88 female calves (22% of T_M) that reached 4 days old were retained as YHs while the rest were culled (at 4 days old). To adjust for natural mortality, 90.5% animals were assumed to survive through weaning (at 70 days old) (E. Cuttance et al., 2017). YHs were moved to an off-site grazing location on May 1st when they were approximately 9 months old and returned on May 1st of the following year when they were approximately 21 ~ 22 months old. We assumed that only the pregnant BHs were returned and mixed with the MA group, while the rest were culled.

Puberty in heifers was assumed to start at 380 days of age (McNaughton et al., 2002) with oestrus cycles occurring every 18 to 24 days thereafter until the animal became pregnant. When a pregnant animal calved or aborted, the next oestrus was assumed to occur after 15 to 49 days (Olds & Seath, 1953) and was assumed to be silent. For simplicity, the PSM for both the MA and BH herds started on October 24th and the mating period lasted for 12 weeks. During the first 6 weeks, animals in oestrus were artificially inseminated. The probability of oestrus detection was set at 78.6% (Anonymous, 2018a) and the probability of conception to artificial insemination was set at 60.0%. For the last 6 weeks, animals in oestrus were naturally mated with breeding bulls with conception probability of 45.0%. This resulted in 66.1% of cows artificially bred during the first 6 weeks being pregnant and 15.0% of cows remaining non-pregnant after 12 weeks of mating period (Anonymous, 2018a).

The gestation period was set at 281 days. A natural abortion rate of 0.00013 per day was applied over the whole gestation period to achieve an overall abortion probability of 3.5% (Weston et al., 2012). Pregnancy scanning was conducted on February 28th for MAs only and all non-pregnant cows were culled at the start day of the drying-off period (May 1st). We assumed that all returning BHs that were pregnant joined the MA herd on May 1st, and we kept the MA herd size constant by (1) selling x number of pregnant MAs older than 6 years where x was the surplus of MAs compared to T_M , or (2) purchasing x number of replacement heifers where x was the shortage of MAs compared to T_M . We assumed that any purchased heifers were susceptible to BVDV. The herd of breeding bulls was not considered in the simulation model since it is common practice for bulls to be present on farm only for the duration of the mating period and for most bulls to be tested for and vaccinated against BVDV (Han, Holter, et al., 2018). An illustration of demographic events for the dairy model is provided in Figure 5.1, and the parameters for transition between demographic groups are described in Table 5.1.

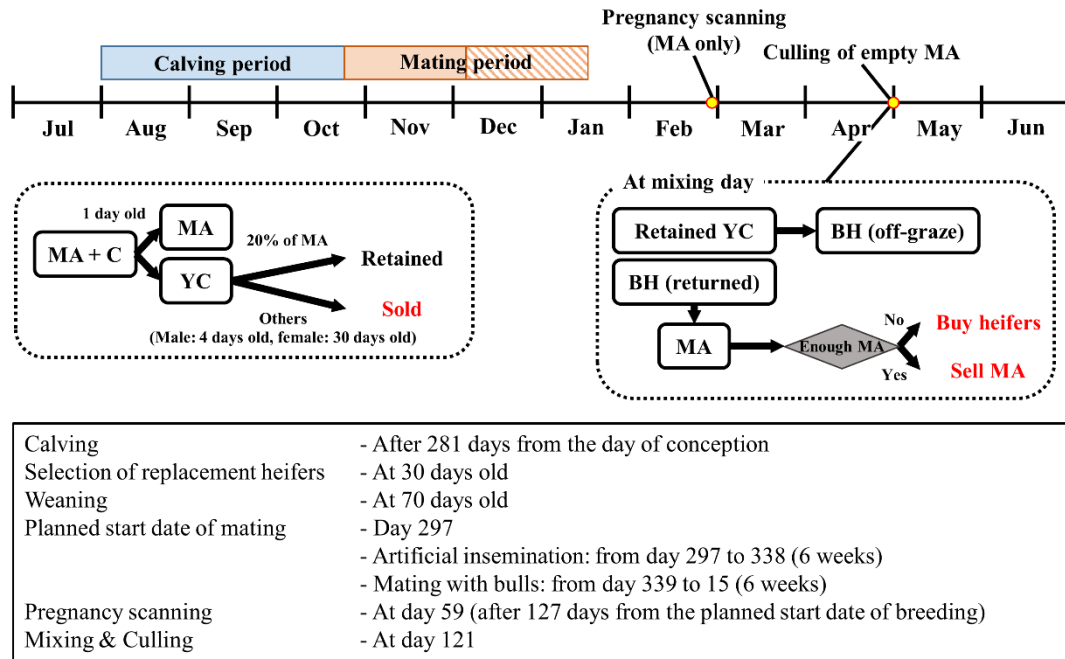


Figure 5. 1. Schematic representation of demographic events in a dairy farm. C, YH, BH, and MA indicate calf herd, young replacement heifer herd, breeding replacement heifer herd, and mixed-aged cow herd, respectively. Filled and shaded area in the mating period indicate the period of artificial insemination and natural mating with bulls, respectively.

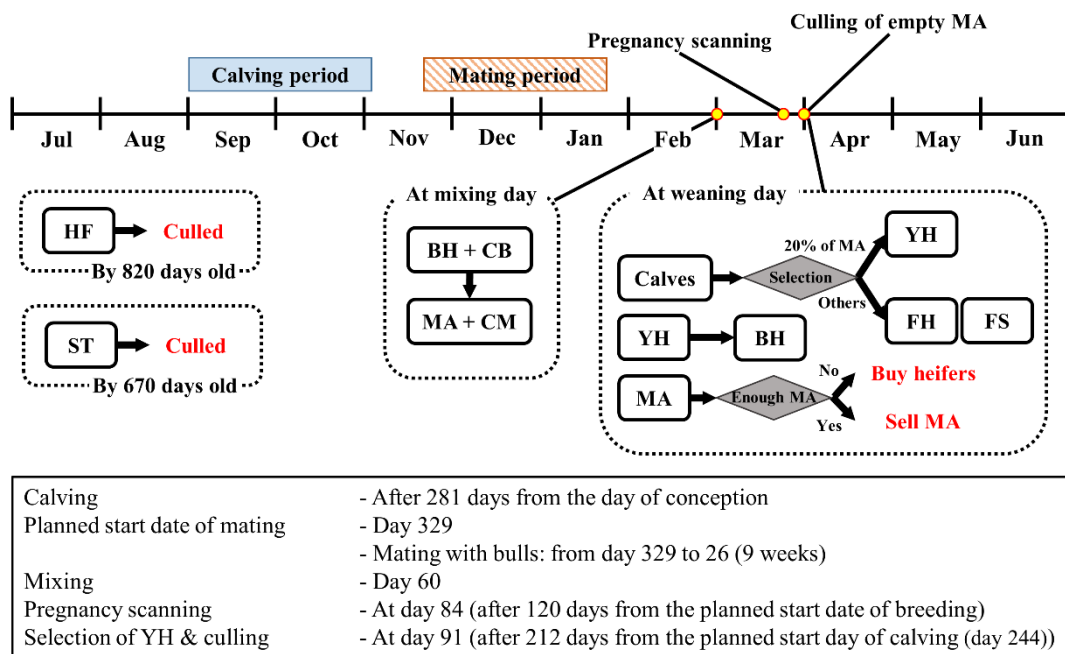


Figure 5. 2. Schematic representation of demographic events in a beef farm. CH, CM, YH, BH, MA, FH, and FS indicate calf herd from breeding heifers, calf herd from mixed-aged cow herd, young heifer herd, breeding heifer herd, mixed-aged cow herd, fattening heifer herd, and fattening steer herd, respectively.

Table 5. 1. Model parameters for demographic and BVDV components.

Parameters	Value (unit)	Reference
Demographic component		
Natural abortion	0.00013 (/day)	Weston et al., (2012)
Natural mortality by weaning:		E. Cuttance et al., (2017); Hickson et al., (2012); C. A. Morris et al., (1986)
Dairy calves	0.0014 (/day)	
Heifer-born beef calves	0.0009 (/day)	
Cow-born beef calves	0.0003 (/day)	
Probability of oestrus detection (dairy only)	0.786	Anonymous, (2018a)
Probability of conception:		Anonymous, (2018a); Geenty & Morris, (2017)
By artificial insemination (dairy only)	0.60	
By natural mating with bulls	D: 0.45 / B: 0.61	
Disease component		
Transmission rates of:		Han et al., (2019); Smith et al., (2014); Viet et al., (2004)
Within-herd by a PI animal (β_P)	D: 0.50 / B: 0.11	
Within-herd by a TI animal (β_T)	$\beta_P \times 0.05$	
Between-herd by a PI animal (β_B)	$\beta_P \times 0.40$	
Between-farm by a PI animal (β_F)	$\beta_P \times 0.40$	
Prevalence of PI animals on neighbouring farms (K_F)	0.02	Assumed
Duration of passive immunity via maternal antibody	$N(155, 31^2)$	Han et al., (2019)
Duration of recovery of TI animals	$U(10, 20)$	Liebler-Tenorio et al., (2004); Müller-Doblies et al., (2004); T. Sandvik et al., (1997)
Mortality of PI animals (μ)	0.0019 (/day)	Ezanno et al., (2007)
Reduced conception of TI and PI animals	0.68	McGowan et al., (1993); Whitmore et al., (1981)
Probability of abortion if:		Grooms et al., (2007); McClurkin et al., (1984); McGowan et al., (1993);
Dams were infected during 0 ~ 41 days of gestation	0.121	Viet et al., (2004); Walz et al., (2017, 2018)

Dams were infected during 42 ~ 150 days of gestation	0.174	
Dams were PI and at 0 ~ 41 days of gestation	0.0086 (/day)	
Dams were PI and at 42 ~150 days of gestation	0.0128 (/day)	
Probability of BVDV infection status of calves if dams were infected during 42 ~ 150 days of gestation:		Viet et al., (2004)
PI calves	0.934	
Passively immunized calves via maternal antibody	0.033	
Recovered calves	0.033	

Key: *PI*, Persistently infected; *TI*, Transiently infected; $N(\cdot, \cdot)$, Normal distribution (mean, variance); $U(\cdot, \cdot)$, Uniform distribution (lower, upper limit); *D*, Dairy; *B*, Beef.

5.3.2. Demographics of beef farms

The typical calving season for beef farms in New Zealand is between August and November with the PSM between early November and early December. More than half of farms have a mating period of 9 weeks or less (McFadden et al., 2005). Calves are weaned and separated from their dams at approximately 6 to 7 months of age at which time a proportion of female calves are selected to become replacement heifers while the others heifer calves and all male calves are raised or sold for finishing. Generally, heifers deliver their first calf at approximately 24 months of age, but some farmers breed heifers for the first time at approximately 26 to 27 months of age with first calving at approximately 36 months of age (McFadden et al., 2005). Most farmers mix the replacement heifers with the MA herd after the heifers' first calving (Hickson et al., 2012).

For the simulation model, we divided the beef population into 7 management groups: Fattening heifers (FH) and steers (FS) for finishing, young replacement heifers (YH: heifers in this herd are bred for the first time), breeding replacement heifers (BH: heifers in this herd calve for the first time and are bred for the second time) and their calves (CH), and mixed-age cows (MA) and their calves (CM).

The start of puberty, oestrus cycle, gestation period, and natural abortion rate were assumed to be the same as dairy heifers/cows (Hickson et al., 2012). The PSM was November 25th, and cows/heifers were grazed with breeding bulls for 9 weeks of the mating period. For animals in oestrus, the probability of conception by naturally mating with breeding bulls was 62.0% (Geenty & Morris, 2017). Off-site grazing was not considered in the beef model since it is not a common practice in the beef industry.

As beef cattle in New Zealand are infrequently handled compared with dairy cattle, we assumed that most of the transitions between demographic management groups occurred on April 1st, which was the day of weaning. We also assumed that the natural mortality rate of calves until weaning differed between CHs and CMs with 14.3% and 5.0%, respectively

(Hickson et al., 2012; C. A. Morris et al., 1986). When the calves were weaned, up to 33 female calves (22% of T_M) were selected from those who were born during the first 3 weeks of the calving period and transferred to the YH group. Other calves were transferred to either FH or FS depending on their sex, and slaughtered when they reached 820 or 670 days of age, respectively (Geenty & Morris, 2017). Transition of YHs into BHs also occurred on the day of weaning, and it was assumed that all BHs joined the MA herd on March 1st which was between the end of the mating period and pregnancy scanning. On March 25th (one week before the day of weaning), pregnancy scanning of MAs was performed and any empty MAs were tagged. We kept the herd size constant by either (1) selling some of the pregnant (i.e. non-tagged) MAs older than 6 years if the number of MAs (= non-tagged MAs + joined BHs) was larger than T_M , or (2) purchasing (BVDV susceptible) replacement heifers if the number of MAs (= non-tagged MAs + joined BHs) was lower than T_M . As in the dairy model, the demographic group of breeding bulls was not included in the beef model. Figure 5.2 illustrates the demographic events of the beef model, and the parameters for transition between demographic groups are described in Table 5.1.

5.3.3. Disease component

For the disease simulation component, the BVDV infection status of individual animals was categorised into the following five mutually exclusive disease states: passively immunized via ingestion of maternal antibodies through colostrum (M), susceptible (S), transiently infected (TI), persistently infected (PI), or recovered (R). In the simulation, the probability of BVDV infection for each susceptible animal (P_{inf}) was determined by the numbers of PI and TI animals within the same group, PI animals in different groups on the same farm, and PI animals on neighbouring farms which could be a source of infection. The occurrence of BVDV infection in an individual animal was determined with a random

Bernoulli event, with the probability of a susceptible individual in a group i being infected on a given day was

$$P_{inf} = 1 - e^{(-\lambda_i)}$$

$$\lambda_i = \beta_P \frac{PI_i}{N_i} + \beta_T \frac{TI_i}{N_i} + \beta_B \sum_j^n \frac{PI_j}{N_i N_j} + \beta_F \frac{K_F}{N_i}, (i \neq j)$$

where λ was the force of BVDV infection in group i at a given day, β_P and β_T were within-group transmission rates by a PI and TI animal, respectively, β_B and β_F were between-group and between-farm transmission rates by a PI animal, respectively, N_i , PI_i , and TI_i were the group size, the number of PI and TI animals in group i on a given day, respectively, and n was the number of groups on a farm on a given day. Given the endemic situation of BVDV, we also considered the local transmission of BVDV by incorporating the overall prevalence of PI animals on the neighbouring farms (K_F) to the force of BVDV infection. We assumed that β_T , β_B , and β_F were proportionally lower than β_P based on our estimation and other published studies (Smith et al., 2014; Viet et al., 2007). Dams and their calves were assumed to be one management group from calving to weaning in the beef model for the purpose of modelling BVDV transmission. Also, since animals in the BH herd in the dairy model were grazed at an off-site location, any between-group transmission to and from these animals (except from neighbouring farms) was not considered in the dairy model.

TI animals recovered and became seropositive after x days of infection, where x was randomly chosen from a uniform distribution between 10 and 20 days (Liebler-Tenorio et al., 2004; Müller-Doblies et al., 2004; T. Sandvik et al., 1997). Once recovered, animals were assumed to have developed life-long immunity against BVDV. We assumed the daily mortality rate of 0.0006/day for PI animals (μ), which indicated an annual survival probability of 50.0% (Ezanno et al., 2007).

We assumed that the duration of protection against BVDV infection via maternal antibodies followed a normal distribution with mean and standard deviation of 155 and 31 days, respectively (Han et al., 2019). Several studies have shown that BVDV can still replicate and induce cell-mediated immune responses in calves with maternal antibodies without causing them to show clinical signs, and that these calves develop stronger immune responses (i.e. an anamnestic response) when they were re-exposed to the virus following the depletion of maternal antibodies (if the initial exposure to the virus was after approximately 14 days old) (Platt et al., 2009; Ridpath et al., 2003). To incorporate the anamnestic immune response of passively immunised calves, we assumed that calves in M status (only over 14 days old) transferred to R status with the probability of P_{inf} per day.

It is well-known that BVDV decreases the fertility of infected animals, and we parameterised that the probability of conception (both by artificial insemination and by mating with bulls) was 0.68 times that of normal animals in both PI and TI animals (McGowan et al., 1993; Whitmore et al., 1981). We also considered the probability of BVDV-induced embryonic loss or abortion, which varied by the stage of gestation at the time of BVDV infection. The probability of embryonic loss was assumed to be 12.1% if animals were infected with BVDV during early gestation (0 ~ 41 days) (McGowan et al., 1993). If BVDV infection occurred during mid-gestation (42 ~ 150 days), the probability of abortion was 17.4% (Grooms et al., 2007; McClurkin et al., 1984; Viet et al., 2004; Walz et al., 2017, 2018). Finally, we assumed that no abortion occurred if animals were infected with BVDV after 150 days of gestation. For simplicity, any TI animals that aborted were assumed to abort on the day of infection. However, the abortion of PI animals was applied as daily rates during the gestation periods (0 ~ 41 days: 0.86%/day, 42 ~ 150 days: 1.28%/day) given the possible persistence of BVDV in uterine tissue (Liebler-Tenorio, 2005).

If animals did not abort, the infection status of their calf was determined by the stage of gestation when the dams were infected. We assumed that a PI dam always produced a PI

calf, and all calves were M or R if their dams were infected with BVDV during 0 ~ 41 or > 150 days of gestation, respectively (Viet et al., 2004). For the calves from dams infected between 42 and 150 days of gestation, the probability of becoming PI was 93.4%, and the rest became either M or R (3.3% for each) (Viet et al., 2004). Schematic representations of transition between BVDV infection status and the BVDV infection status of new born calves are illustrated in Figure 5.3. Parameters related to BVDV infection are provided in Table 5.1.

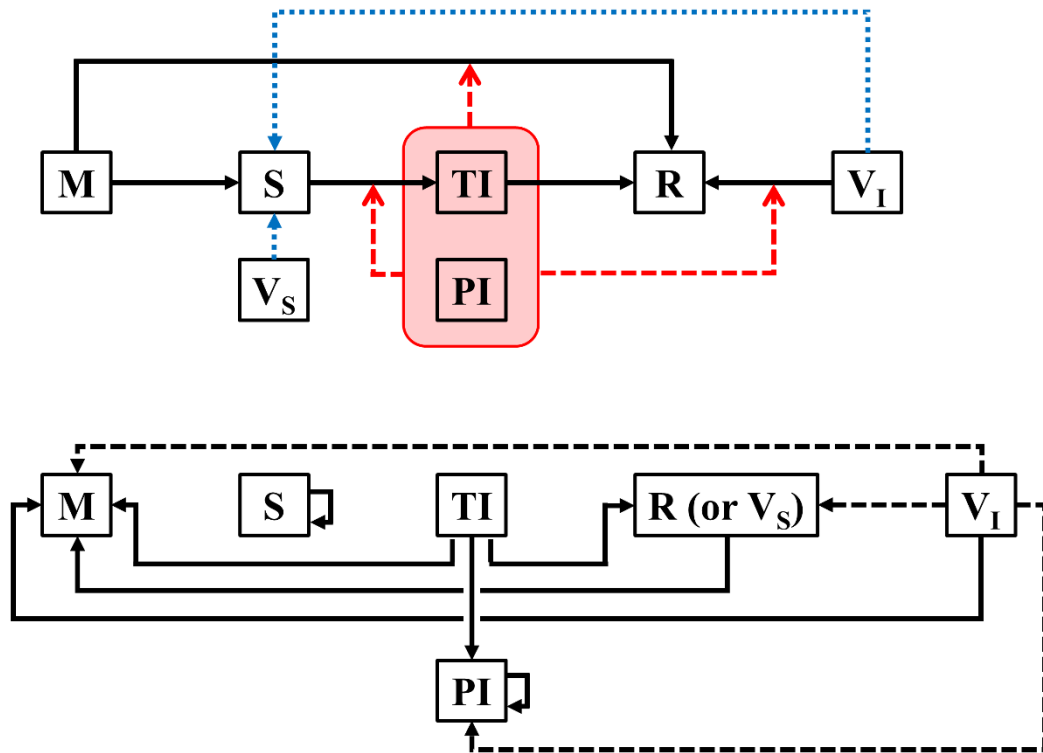


Figure 5. 3. Schematic representation of transition between BVDV infection status (top) and BVDV infection status of new born calves (bottom). BVDV infection status consists of passively immunised via maternal antibody (M), susceptible (S), transiently infected (TI), persistently infected (PI), recovered (R), vaccinated insufficiently to prevent fetal infection (V_I), and vaccinated sufficiently to prevent fetal infection (V_S). Top: Solid black arrows indicate the transition between BVDV infection status, blue dotted arrows indicate the decay of immunity conferred by vaccination, and red dashed arrows illustrate the force of BVDV infection. Bottom: Solid arrows indicate the BVDV infection status of calves depending on the infection status of their dams. Dashed arrows from V_I illustrate the BVDV infection status of calves if V_I dams were infected.

Table 5. 2. Model parameters for production/economic and control components.

Parameters	Value (unit)	Reference
Production and economic component		
Production of milk solid of:		Anonymous, (2018a)
Primiparous cows	1.33 (kg/day)	
Multiparous cows	1.76 (kg/day)	
Reduced production of milk solid in PI cows	0.478	Voges et al., (2006)
Reduced live-weight gain per day of:		Moffat & Bruere, (2018); Voges et al., (2006)
PI animals	0.81	
TI animals	0.92	
Price of milk solid	6.79 (NZ\$/kg)	Anonymous, (2018a)
Price of cattle:		Anonymous, (2019b)
Bobby calf	40.00 (NZ\$)	
Pregnant mixed-age cow	1,700.00 (NZ\$)	
Replacement heifer	1,300.00 (NZ\$)	
Price of carcass of:		Anonymous, (2019c)
Calf or young heifer	3.45 (NZ\$/kg)	
Fattening steer or heifer, and breeding heifer	5.76 (NZ\$/kg)	
Mixed-age cow	4.41 (NZ\$/kg)	
Cost of artificial insemination	31.85 (NZ\$)	Anonymous, (2019a)
Cost of treatment	27.00 (NZ\$)	Our estimation
Cost of BVDV RT-PCR test on bulk tank milk	170.90 (NZ\$)	Anonymous, (2019d)
Cost of BVDV antigen ELISA	17.09 (NZ\$) *	Reichel et al., (2008)
Cost of BVDV vaccination	6.74 (NZ\$) *	Reichel et al., (2008)
Cost of double fencing	1.80 (NZ\$/m) †	Anonymous, (2016d)

Control component		
Test performance:		Lanyon et al., (2013)
Sensitivity of BVDV RT-PCR test	0.999	
Specificity of BVDV RT-PCR test	0.999	
Sensitivity of BVDV antigen ELISA	0.835	
Specificity of BVDV antigen ELISA	0.994	
Probability of vaccination outcome (after second dose)		Estimated
Vaccine failure	0.193	(see Section 3 of Appendix 3)
Insufficient fetal protection (V_I)	0.735	
Sufficient fetal protection (V_S)	0.072	
Probability of abortion if:		Estimated
V_I dams were infected during 0 ~ 41 days of gestation	0.077	(see Section 3 of Appendix 3)
V_I dams were infected during 42 ~ 150 days of gestation	0.076	
Probability of BVDV infection status of calves if V_I dams were infected during 42 ~ 150 days of gestation:		Estimated (see Section 3 of Appendix 3)
PI calves	0.232	
Passively immunised calves via maternal antibody	0.559	
Recovered calves	0.209	

* Monetary values were standardised using 1.9% inflation rate from the year of original publication (NZ\$ 7.90 was added to the cost of ELISA for veterinary service).

† The distance of double fencing was estimated based on the total herd size with the circular farm area of 2.84 hectare/cattle. It resulted the cost of installing double fencing for both dairy and beef farms as NZ\$ 4,608.52.

5.3.4. Production and economic component

In this study, we considered only the direct production impacts of BVDV infection on milk yields, growth rates, reproductive performance, and mortality. In order to estimate the annual gross farm revenue (GFR) for a dairy farm, total milk solids (MS) produced in each production year (from PSC to the next PSC) was calculated. In the model, all MAs started lactation from the day of calving and were dried-off on May 1st. We assumed that the average MS produced per day for primiparous and multiparous cows was 1.33 and 1.76 kg/day, respectively (Anonymous, 2018a). To account for the effect of BVDV on milk production, we assumed that PI animals produced 47.8% of the total MS compared with normal (S and R) animals per day (Voges et al., 2006). For simplicity, milk produced by TI animals was assumed to be discarded and cows that aborted at any stage of gestation were assumed not to produce milk in the following production year. Farm-gate revenue (farmer's earning after deducting the cost of milk transfer, manufacturing, etc.) was NZ\$ 6.79 per kg MS (Anonymous, 2018d).

Besides milk production, revenue from selling bobby calves and surplus MAs were also considered by multiplying the number of each type of animals sold per production year and the expected price of the animal (bobby calves: NZ\$ 40.00, MAs: NZ\$ 1,700.00) (Anonymous, 2019b). We also incorporated the revenue earned by culling empty and aborted MAs by calculating the value of carcass (carcass weight \times NZ\$ 4.41/kg) (Anonymous, 2019c). The carcass weight of each individual was assumed to be 50.0% of the live-weight at the time of culling, and live-weight gain per day (LWG/day) of animals was estimated based Von Bertalanffy function (Brown et al., 1976; García-Muñiz et al., 1998; Geenty & Morris, 2017; Hickson et al., 2015; S. Morris et al., 2006). Details of the estimation of LWG/day is provided in Section 2 of Appendix 3. To adjust the impact of BVDV infection on LWG/day, we assumed that PI and TI animals had 0.81 and 0.92 time of LWG/day of normal (M, S and R) animals (Moffat & Bruere, 2018; Voges et al., 2006). In cases of BVDV induced abortion for an animal that had conceived by artificial

insemination, the cost of artificial insemination (NZ\$ 31.85) was deducted from GFR (Anonymous, 2019a). We also deducted the cost of purchasing replacement heifers (NZ\$ 1,300.00/heifer) and treatment for TI animals (lactating MA only) (NZ\$ 27.00/TI animal) from GFR (Anonymous, 2019b). For the beef model, we assumed that GFR for each production year was only affected by the overall carcass weight of culled and slaughtered animals (FH and FS: NZ\$ 5.76/kg, MA: NZ\$ 4.41/kg) and the number of surplus MAs sold, however, we did not consider the treatment of TI animals in the beef model (Anonymous, 2019c). Parameters related to production and economic components are provided in Table 5.2.

5.3.5. BVDV control strategies

In this study, we examined different BVDV control strategies as a combination of different control options in three categories: (1) annual BVDV testing and culling of identified PI animals, (2) vaccination of breeding animals, and (3) increasing biosecurity by double fencing. We assumed that a BVDV control strategy was continually performed during the simulation once it was implemented.

For annual BVDV testing and culling, we considered two sub-options (i) testing all female calves selected as replacement heifers using BVDV antigen (Ag) ELISA on ear notch sample, followed by culling of confirmed PI animals, or (ii) screening of BVDV using RT-PCR on bulk tank milk (BTM) for the dairy farm or using individual BVDV Ag ELISA on serum samples from randomly chosen 15 YHs for the beef farm, followed by a PI hunt (i.e. testing all breeding animals using BVDV Ag ELISA with culling of any test-positive animals) of all breeding animals if the screening test was positive. It was assumed that all female calves were tested (i.e. initial test) at 4 days old for dairy calves and on the day of weaning for beef calves. All test positive calves were re-tested (i.e. subsequent test) 28 days after the initial test to confirm their infection status as PI, and all confirmed PI

calves were culled immediately. When the subsequent test was conducted, the dam of the tested calves was also sampled and tested using the Ag ELISA and culled immediately if the result was positive. For annual screening, RT-PCR on BTM for the dairy farm was conducted at the PSM, and sampling of the 15 YHs on a beef farm occurred on March 1st. Similar to the testing calves, a PI hunt was conducted 28 days after the screening test if RT-PCR of BTM (Ag ELISA of at least one animal among 15 YHs) was test positive for a dairy (beef) farm. When animals were culled due to the test positivity, the carcass value of the animals were added to GFR.

For vaccination, we considered the use of killed-vaccines as they are the only type available in New Zealand. We assumed that (1) the duration of protection after the second dose of BVDV vaccination was 6 months, whereas the duration after the third dose (and thereafter) was 12 months, and (2) BVDV vaccination after the second dose (and thereafter) resulted in three outcomes (i) not enough immunity to prevent clinical diseases (vaccine failure: 19.3%), (ii) enough immunity to prevent clinical disease but insufficient to guarantee the prevention of fetal infection (V_I , insufficient fetal protection: 73.5%), and (iii) sufficient immunity to prevent fetal infection (V_S , sufficient fetal protection: 7.2%) (Downey-Slinker et al., 2016; McArthur, 2004; Newcomer et al., 2015; Packianathan et al., 2017; Ridpath, 2013). The efficacy of BVDV vaccination of V_I animals against fetal infection was assumed to vary according to the stage of gestation at the time of infection (Table 5.2). We also assumed that the first vaccination had no protective effect and there was no effect of BVDV vaccination on TI or PI animals. Detailed information about the parameterisation of BVDV vaccine efficacy is provided in Appendix 3 (see Section 3), and the graphical illustration of reproductive outcomes following BVDV infection on V_I dams is provided in Figure 5.3.

As the duration of BVDV protection was dependent on the number of vaccine doses given, we considered four vaccination programmes that administering to breeding animals: a programme that administering three doses before the first mating (V_1), and a programme

that administering two doses before the first mating (V2), each of which could involve annual boosting (V1AB, V2AB). For the V1 of the dairy model, all replacement heifers in the YH herd were vaccinated for the first time on April 1st and the second time on May 1st to immunise animals before being grazed on off-site. The third dose was administered on September 26th which was 4 weeks before the PSM. For the V2 of the dairy model, all animals in the BH herd were first vaccinated on August 31st and the second time on September 26th. Dairy cows in the MA herd were assumed to be vaccinated on September 26th for annual boosting for V1AB and V2AB. For the V1 of the beef model, the first vaccination was administered to calves on the day of weaning. The second and third doses were given on May 1st and October 29th (4 weeks before the PSM), respectively, to YHs. For the V2 programme, the first and second doses of the beef farm were administered to YHs on October 1st and 29th, respectively. The annual booster was administered on October 29th to all breeding females in MA and BH herds for both V1AB and V2AB. The cost of BVDV vaccination was NZ\$ 6.74 per animal per vaccination (Reichel et al., 2008). Table 5.2 provides model parameters related to control measures.

We assumed that double fencing totally prevented the transmission of BVDV from neighbouring farms. The cost for installing double fencing varied by farm, and was calculated at NZ\$ 1.80/metre (Anonymous, 2016d) where the fencing distance was estimated based on the total herd size with the circular farm area of 2.84 hectare/cattle (Anonymous, 2018a). The cost of installing double fencing only applied in the first year of simulation, and 10% of the cost was assumed to be spent in following years for maintenance.

5.3.6. Model initiation and BVDV outbreak

We initiated the dairy model with 400 pregnant MAs and 80 BHs. In the beef model, the initial population consisted of 150 pregnant MAs, 30 pregnant BHs, and 30 YHs. The simulations started on the day of the PSC, and the occurrence of demographic events (e.g.

natural mortality, conception, abortion, oestrous detection) was determined based on conducting Bernoulli trials with the relevant probability for each individual on each day of the simulation. We first ran the models for 5 years as a burn-in period to stabilise the population demography, then introduced female PI calves into the farm at the beginning of year 6 to mimic BVDV outbreak. PI calves were introduced by converting x number of susceptible new born female calves into PI, where x was 4.0% of $T_M(\rho)$ (Thompson, 2005). When PI calves were introduced, their dam's infection status was also converted to R (recovered) as it was assumed that the BVDV outbreak was due to purchasing Trojan cows (i.e. a dam carrying a PI fetus). Local transmission was also allowed from the beginning of year 6 until the end of simulation by assuming the overall prevalence of PI animals on neighbouring farms (K_F) as 2.0%. We ran the simulation model for an additional 5 years to monitor BVDV dynamics and the economic impacts of applying different BVDV control strategies. The simulation period of 5 years was chosen to estimate the economic impacts of a single incursion event on BVDV naïve farms and producing a realistic timeframe for farmers to consider the economics of investing in BVDV control.

5.3.7. BVDV dynamics and economic analysis

First, we simulated the models both with and without BVDV (i.e. both ρ and K_F being zero) for 1,000 iterations each while implementing no control measures to estimate the economic impacts of a BVDV outbreak on dairy and beef farms. The output measures for the model were (1) the number of TI and PI animals per year, and (2) the loss of total GFR due to BVDV to estimate the dynamics and economic impact of BVDV in naïve farms. The total GFR was calculated as a sum of annual GFRs during the simulation period while adjusting for 1.9% of discounting rate per year.

Next, we evaluated the economic impact of different control strategies by running 30 different scenarios for both dairy and beef models. Each scenario was a combination of

different test and cull, vaccination, and double fencing options (Table 5.3), and was iterated 1,000 times to adjust for the variation caused by stochasticity, resulting in 30,000 simulations for each of the dairy and beef models. For each simulation, we calculated the total GFR and total cost of BVDV control during the simulation period (i.e. cumulative cost while discounting by 1.9% per year). We then established a multivariable linear regression model with the total GFR as an outcome and the various BVDV control strategies as explanatory variables. The constant of the multivariable model represented the total GFR without any control measures, and the coefficients were the average benefits (i.e. surplus in total GFR) by adapting relevant control strategies. Coefficients of the model were used to analyse the benefit/cost ratio of BVDV control strategies, which was calculated as a ratio of the coefficient to the total cost of BVDV control strategies. Given the amount of output data, we only considered any variable that had P value < 0.01 as statistically significant in multivariable models. Since the purpose of the regression models was to estimate the economic impact or efficacy of all control strategies, we examined only the full models that considered all the explanatory variables, and therefore did not use any type of model selection process based on statistical significance. The simulation models in this study were coded in the C programming language and were called in R (version 3.5.2) to substitute the varying parameter values (R Core Team, 2018). All analyses of the simulation results were performed in R.

Table 5. 3. List of BVDV control options in three different categories. A BVDV control strategy was a combination of three control options from each of category (30 strategies in total).

Category	Control option
Annual testing and culling of identified PI animals	No
	Annual testing and culling of female breeding calves
	Annual screening of BVDV (BTM) or Ag (serum of 15 young heifers)
Vaccination of female breeding animals	No
	Administering three doses before the first mating
	Administering three doses before the first mating with annual boosters
	Administering two doses before the first mating
Increasing biosecurity by double fencing	Administering two doses before the first mating with annual boosters
	No
	Yes

Key: BTM, bulk tank milk; Ag, antigen.

5.3.8. Sensitivity analysis

Sensitivity analysis was conducted to investigate the impact of different parameter values (β_P , β_B , β_F , K_F , μ , and ρ) on the model outputs (number of TI, PI animals, and total GFR). We used factorial design to examine different values for each parameter, resulting in 4,800 combinations of different parameter values for both dairy and beef models. We ran the simulation 100 times for each combination (960,000 simulations in total), and simulated model outputs were aggregated into a dataset. Each output was then analysed as an outcome of multivariable linear regressions with all parameters as explanatory variables (Saltelli et al., 2019). We examined the coefficients of parameters to measure the change in each output. We also measured the contribution of parameters to the variance of each output as a proportion of the sum of squares related to the parameter to the total sum of squares of the model (Ezanno et al., 2007). Parameter values used for the sensitivity analysis are shown in Table 5.4.

Table 5. 4. Parameter values used for the sensitivity analysis.

Parameters	Examined values
Within-herd BVDV transmission rate by a PI animal (β_P)	Dairy: [0.3, 0.4, 0.5, 0.6, 0.7] Beef: [0.09, 0.10, 0.11, 0.12, 0.13]
Relative proportion of β_B to β_P	[20%, 30%, 40%, 50%]
Relative proportion of β_F to β_P	[20%, 30%, 40%, 50%]
Prevalence of PI animals in the neighbouring farms (K_F)	[1%, 2%, 3%, 4%]
Mortality of PI animals (μ)	[0.0019, 0.0009, 0.0006]*
Number of introduced Trojan cows as a proportion of T_M (ρ)	[2%, 3%, 4%, 5%, 6%]

* Mortality rates of PI animals were based on the half-life of 12, 24, and 36 months, respectively.

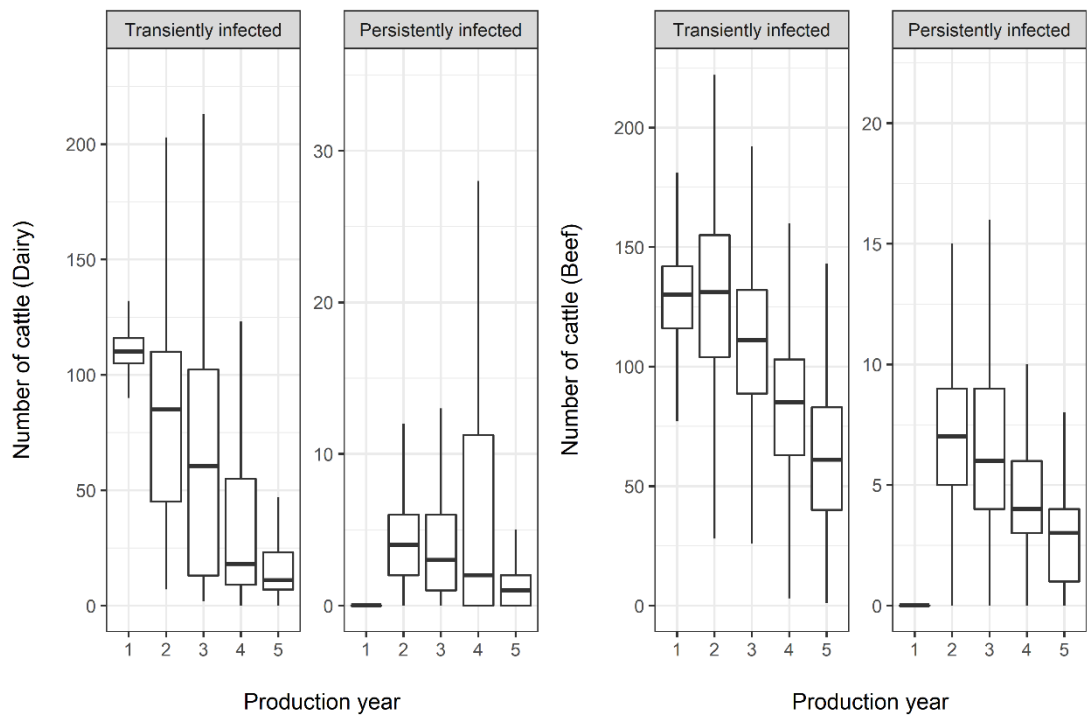


Figure 5. 4. The number of transiently and persistently infected animals during the simulation period in New Zealand dairy (left) and beef (right) farms.

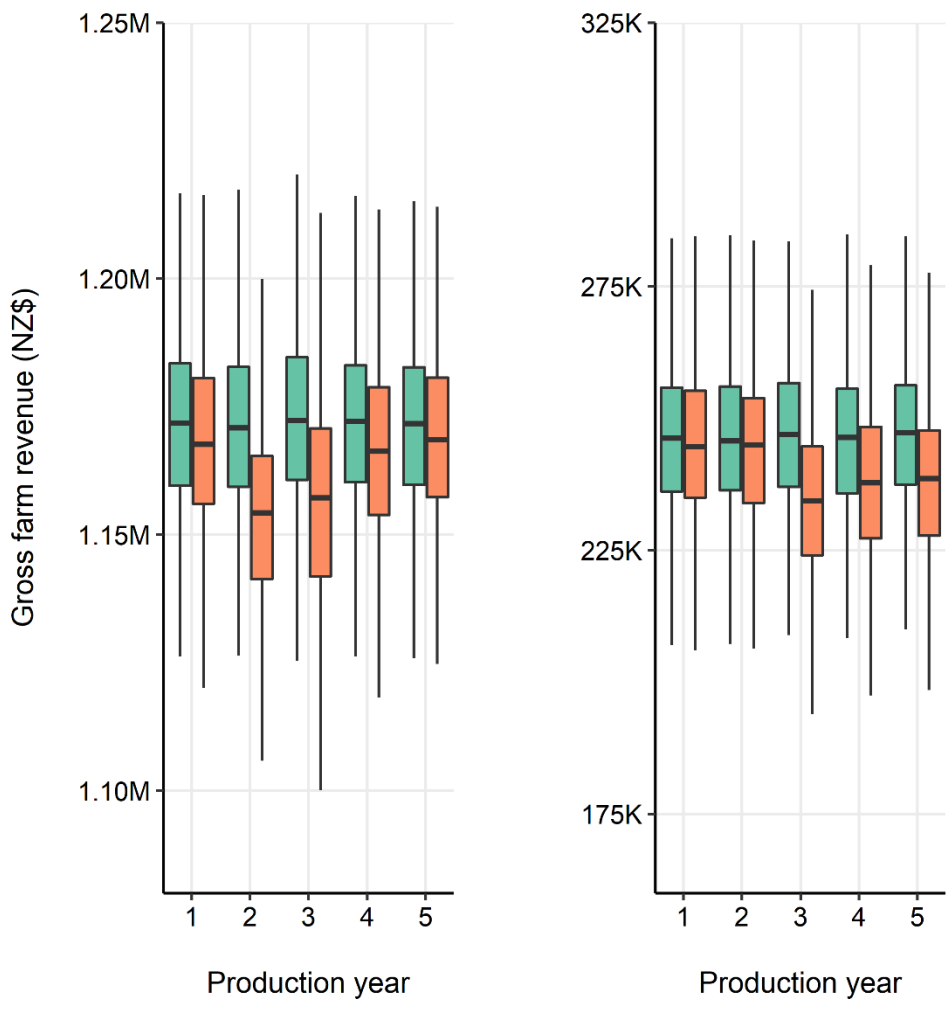


Figure 5. 5. The economic impact of BVDV introduction on naïve dairy (left) and beef (right) farms in New Zealand. Green and orange bars indicate the gross farm revenue without and with BVDV, respectively.

5.4. Results

5.4.1. Dynamics of BVDV and its economic impact

The number of transiently and persistently infected animals on both dairy and beef farms during the 5 years are illustrated in Figure 5.4. On average, the number of PI animals (excluding initially introduced PI calves) on both dairy and beef farms was highest one year after initial introduction. By the end of the simulation period (year 5), BVDV had almost been eliminated naturally on the dairy farm, whereas some TI and PI animals still persisted on the beef farm.

Figure 5.5 illustrates the annual GFR during the simulation period with and without BVDV. The largest reduction in GFR was observed during year 2 ~ 3 on the dairy farm, whereas it was at year 3 on the beef farm. The annual GFR was restored to the level of year 1 at the final year (year 5) on the dairy farm, however, it continued to be affected on the beef farm indicating that the economic impact of BVDV in a beef farm would last longer than 5 years. Mean (standard deviation) losses in total GFR were approximately NZ\$ 44,400 (1,555) and NZ\$ 30,900 (912) on dairy and beef farms, respectively. Dividing the total loss by the number of MAs on each farm, the economic losses per MA (95% confidence interval) were NZ\$ 22.22 (20.69 ~ 23.74) per year for the dairy farm and NZ\$ 41.19 (38.81 ~ 43.58) per year for the beef farm.

5.4.2. Cost-benefit of BVDV control strategies

Table 5.5 and 5.6 show the results of multivariable linear regression models on total GFR, and the cost and benefit of implementing BVDV control are illustrated in Figure 5.6 with benefit/cost ratio. On both types of farm, the highest benefit was incurred when annual testing of breeding calves was implemented with the vaccination programme that administering three doses before the first mating and annual boosters (NZ\$ 22,864.65 for dairy, NZ\$ 21,582.50 for beef farms).

In a dairy farm, only the BVDV control strategies of annual testing of female calves incurred significantly higher benefit compared to its cost (B/C ratio of 2.02). Other strategies, such as the annual testing of calves with double fencing, generated marginally higher benefit than the control cost. In a beef farm, any control strategies that included vaccination with ongoing annual boosters being administered (either V1AB or V2AB) incurred higher benefit than cost as long as double fencing was not considered in the strategies. The highest benefit/cost ratio of 2.02 was observed when only a vaccination programme consisting of two doses before first mating with annual boosting was implemented. Some BVDV control strategies, such as annual screening only, vaccination only without administering annual boosters, or double fencing only, were not economically beneficial on dairy or beef farms based on our models.

Table 5. 5. Benefit and cost of implementing BVDV testing and culling (TC), vaccination (VC) with/without annual boosting (AB), and double fencing (DF) in New Zealand dairy farm. Benefit-cost (B/C) ratios of control strategies with significant benefit are shown in bold.

Control options				Gross farm revenue		Control cost	B/C ratio
TC	VC	AB	DF	Coefficient *	P value		(95% CI interval)
				5,599,708.42	< 0.001		
			✓	3,334.23	0.046	11,461.78	0.29 (0.01, 0.58)
	1			-5,365.87	< 0.001	8,710.92	-0.62 (-0.99, -0.24)
	1		✓	-391.19	0.82	20,186.81	-0.02 (-0.18, 0.14)
	1	✓		16,879.25	< 0.001	26,672.48	0.63 (0.51, 0.76)
	1	✓	✓	15,444.55	< 0.001	38,127.40	0.41 (0.32, 0.49)
	2			-17,045.09	< 0.001	5,062.94	-3.37 (-4.01, -2.72)
	2		✓	-12,612.55	< 0.001	16,527.15	-0.76 (-0.96, -0.57)
	2	✓		319.83	0.85	22,971.96	0.01 (-0.13, 0.16)
	2	✓	✓	3,013.12	0.07	34,433.93	0.09 (-0.01, 0.18)
C				15,839.51	< 0.001	7,837.05	2.02 (1.60, 2.44)
C			✓	19,558.63	< 0.001	19,248.79	1.02 (0.85, 1.19)
C	1			15,264.38	< 0.001	16,430.06	0.93 (0.73, 1.13)
C	1		✓	18,922.61	< 0.001	27,870.03	0.68 (0.56, 0.80)
C	1	✓		22,620.82	< 0.001	34,325.46	0.66 (0.56, 0.75)
C	1	✓	✓	22,864.65	< 0.001	45,777.77	0.50 (0.43, 0.57)
C	2			7,680.56	< 0.001	12,867.75	0.60 (0.34, 0.85)
C	2		✓	13,753.00	< 0.001	24,299.27	0.57 (0.43, 0.70)
C	2	✓		18,158.10	< 0.001	30,728.76	0.59 (0.48, 0.70)
C	2	✓	✓	19,994.54	< 0.001	42,188.03	0.47 (0.40, 0.55)
S				-2,596.41	0.12	12,781.46	-0.20 (-0.46, 0.05)
S			✓	5,871.85	< 0.001	20,759.68	0.28 (0.13, 0.44)
S	1			-2,772.11	0.10	21,964.89	-0.13 (-0.28, 0.02)
S	1		✓	4,995.77	< 0.01	29,716.38	0.17 (0.06, 0.28)
S	1	✓		12,087.36	< 0.001	32,296.11	0.37 (0.27, 0.48)
S	1	✓	✓	15,398.28	< 0.001	42,608.76	0.36 (0.28, 0.44)
S	2			-10,988.67	< 0.001	18,612.09	-0.59 (-0.77, -0.41)
S	2		✓	-7,466.35	< 0.001	27,265.51	-0.27 (-0.39, -0.15)
S	2	✓		413.32	0.80	28,806.18	0.01 (-0.10, 0.13)
S	2	✓	✓	2,339.74	0.16	39,669.77	0.06 (-0.02, 0.14)

Key: C, annual testing of breeding calves; S, annual screening test; 1, administering three vaccine doses before the first mating; 2, administering two vaccine doses before the first mating; CI, confidence interval.

* Standard error of intercept and coefficients were 1,179.34 and 1,667.83, respectively.

Table 5. 6. Benefit and cost of implementing BVDV testing and culling (TC), vaccination (VC) with/without annual boosting (AB), and double fencing (DF) in New Zealand dairy farm. Benefit-cost (B/C) ratios of control strategies with significant benefit are shown in bold.

Control options				Gross farm revenue		Control cost	B/C ratio
TC	VC	AB	DF	Coefficient *	P value		(95% CI interval)
				1,155,505.96	< 0.001		
			✓	1,625.02	0.07	11,461.78	0.14 (-0.01, 0.29)
	1			1,650.24	0.06	3,653.23	0.45 (-0.02, 0.93)
	1		✓	1,507.78	0.09	15,115.14	0.1 (-0.02, 0.21)
	1	✓		20,112.43	< 0.001	12,065.61	1.67 (1.52, 1.81)
	1	✓	✓	21,060.75	< 0.001	23,527.52	0.9 (0.82, 0.97)
	2			946.39	0.29	2,131.42	0.44 (-0.37, 1.26)
	2		✓	2,462.03	< 0.01	13,593.09	0.18 (0.05, 0.31)
	2	✓		18,261.73	< 0.001	9,031.74	2.02 (1.83, 2.21)
	2	✓	✓	17,889.89	< 0.001	20,493.66	0.87 (0.79, 0.96)
C				1,782.37	0.05	2,823.56	0.63 (0.01, 1.25)
C			✓	1,526.92	0.09	14,289.19	0.11 (-0.02, 0.23)
C	1			2,329.24	< 0.01	6,463.88	0.36 (0.09, 0.63)
C	1		✓	2,202.30	0.01	17,926.57	0.12 (0.03, 0.22)
C	1	✓		21,471.95	< 0.001	14,813.76	1.45 (1.33, 1.57)
C	1	✓	✓	21,582.50	< 0.001	26,271.46	0.82 (0.76, 0.89)
C	2			2,996.56	< 0.01	4,947.30	0.61 (0.25, 0.96)
C	2		✓	3,454.11	< 0.001	16,409.54	0.21 (0.1, 0.32)
C	2	✓		18,930.05	< 0.001	11,789.22	1.61 (1.46, 1.75)
C	2	✓	✓	19,252.41	< 0.001	23,249.34	0.83 (0.75, 0.9)
S				1,025.18	0.25	4,850.98	0.21 (-0.15, 0.57)
S			✓	590.66	0.51	15,998.81	0.04 (-0.07, 0.15)
S	1			800.95	0.37	7,975.19	0.1 (-0.12, 0.32)
S	1		✓	1,387.72	0.12	19,393.08	0.07 (-0.02, 0.16)
S	1	✓		20,927.91	< 0.001	15,194.02	1.38 (1.26, 1.49)
S	1	✓	✓	20,624.91	< 0.001	26,722.75	0.77 (0.71, 0.84)
S	2			854.89	0.34	7,101.91	0.12 (-0.12, 0.37)
S	2		✓	1,415.76	0.11	18,524.80	0.08 (-0.02, 0.17)
S	2	✓		18,210.25	< 0.001	12,791.07	1.42 (1.29, 1.56)
S	2	✓	✓	17,917.75	< 0.001	24,244.34	0.74 (0.67, 0.81)

Key: C, annual testing of breeding calves; S, annual screening test; 1, administering three vaccine doses before the first mating; 2, administering two vaccine doses before the first mating; CI, confidence interval.

* Standard error of intercept and coefficients were 1,179.34 and 1,667.83, respectively.

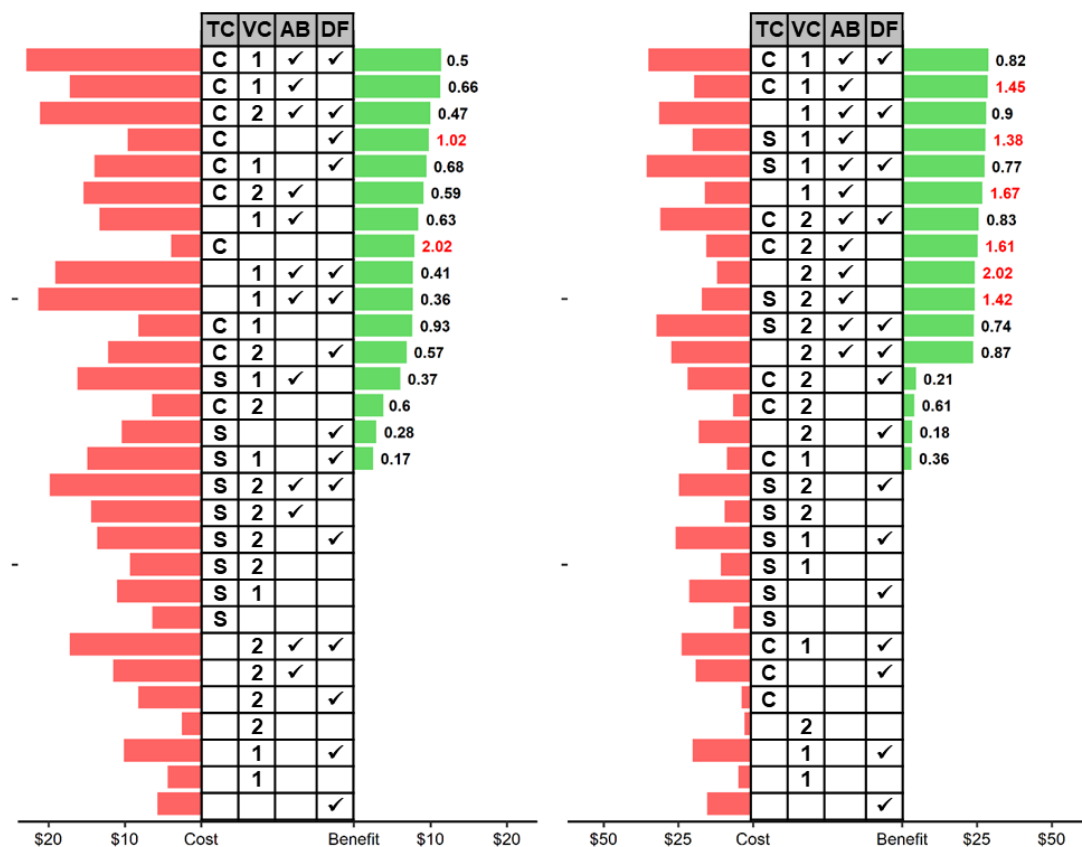


Figure 5. 6. Average annual benefit (green) and cost (red) of BVDV control strategies per mixed-age cow for dairy (left) and beef (right) farms in New Zealand. Non-significant or negative benefits are not presented in this figure, and any average benefit/cost ratio of >1 is written in red. TC, VC, AB, and DF indicate testing and culling, vaccination, annual boosting, and double fencing, respectively. C, S, 1 and 2 indicate the annual testing of female breeding calves, annual screening test, administering three vaccine doses before the first mating, and administering two vaccine doses before the first mating, respectively.

5.4.3. Sensitivity analysis

For the dairy farm, the largest variation in the number of TI animals was explained by the within-herd BVDV transmission rate by a PI animal (43.4%), followed by mortality of PI animals (25.6%) and number of introduced Trojan cows (21.6%). The number of PI animals was mostly affected by the mortality of PI animals (35.3%), number of introduced Trojan cows (31.7%), and within-herd BVDV transmission rate by a PI animal (24.6%). For the beef farm, the variation of both the number of TI and PI animals was mainly explained by the mortality of PI animals (70.3% and 68.5%, respectively). The decrease in GFR due to BVDV infection was dominated by the number of introduced Trojan cows in both dairy and beef farm (66.4% and 66.8%, respectively). Detailed results from the sensitivity analysis are provided in Section 4 of Appendix 3.

5.5. Discussion

In this study, we estimated the economic impacts of BVDV in New Zealand dairy and beef pastoral farming systems using an individual-based modelling approach. We used an individual-based modelling approach because the demographic characteristics of individual animals (e.g. age, sex, and production type) infected with BVDV could have a significant impact on the magnitude of production losses (Richter et al., 2017). Furthermore, New Zealand cattle farming systems are highly seasonal with management events occurring at single, discrete time points, rather than continuously across the year, so it is more difficult to capture by using a traditional compartmental modelling approach. Using an individual-animal based simulation approach, we believe the direct economic losses due to BVDV for typical dairy and beef farms in New Zealand were properly estimated.

Based on this study, the total direct losses caused by BVDV outbreak on the dairy farm was NZ\$ 22.22 per dairy MA per year and NZ\$ 41.19 per beef MA per year. Interestingly, the total direct losses per cow due to BVDV in other countries have been reported to be

generally lower in the beef industry compared to the dairy industry (Richter et al., 2017). This discordant result from our study compared to other countries was possibly due to the unique management characteristics in the New Zealand dairy industry; a narrow seasonal breeding period which limits the risk period for generating PI calves, selling/culling most calves at approximately 4 days old, grazing heifers off-site, and lower average milk yields per cow. We suspect that the early removal of new born PI calves from the property and the limited within-farm transmission due to off-site grazing of replacement heifers likely protect against more severe BVDV outbreaks. Consequently, the reduced infectious pressure in this production type would have limited impact on reproduction and lactation, which are the main contributors to economic losses in dairy farms. Although it suggests that the relative economic impact of BVDV in the New Zealand dairy industry might not be as strong as in other countries, it should be noted that the total direct losses will vary between farms depending on the infectious burden of BVDV (e.g. introduced number of PI animals) or the time or duration of BVDV introduction, as well as other farm management factors, such as the duration of the mating period or the proportion of MAs being replaced (Damman et al., 2015).

Several modelling studies have been conducted to estimate the economic losses of BVDV on New Zealand dairy farms. Previously, Heuer et al., (2007) estimated that the economic losses due to BVDV was NZ\$ 109.05 per cow per year (adjusted for inflation rate since original publication). In a long-term period, Reichel et al., (2008) calculated the total loss of BVDV infection as NZ\$ 43.28 per cow per year (adjusted for inflation rate since original publication). Our estimation of total direct loss per dairy cow per year (NZ\$ 22.22) was lower than these previously estimated values, however, we believe the previously reported values would be overestimated as the previous studies were based on the average performance of BVDV infected herds, where the low performance of BVDV infected herds would not be entirely contributed to BVDV infection (Heuer et al., 2007; Reichel et al., 2008). Compared to those studies, our results would be more conservative

estimates as we considered the direct impact of BVDV infection on the production of each individual animal. However, it is worth noting that the estimated total losses in this study did not incorporate other indirect impacts, such as increased susceptibility to other diseases among infected animals.

To the best of our knowledge, no studies have been conducted to estimate the direct losses due to a BVDV outbreak on New Zealand beef farms. This is mainly due to the logistic challenges of handling extensively grazed herds, which makes it difficult to monitor production (e.g. live weight gain per day) and the infection status of individual animals on a regular basis (Han, Weir, et al., 2018). Therefore, we used a disease simulation approach to estimate the economic impact of BVDV outbreak on New Zealand beef farms assuming BVDV infection would affect the live-weight gain per day of infected animals. However, the estimated loss based on BVDV simulation models highly depends on the modelling assumptions, such as demographic structure or parameter values (Damman et al., 2015). For example, Gunn et al., (2004) measured the annual economic loss due to BVDV outbreak in Scottish beef herds as NZ\$ 129.96 per cow (adjusting for inflation rate from the year of original publication), however, this figure was calculated by counting the number of TI or PI animals generated while ignoring the BVDV status of other animals (i.e. whether they were susceptible, immune, or recovered). A more recent study modelled the decrease of production due to BVDV outbreak in French beef herds (Damman et al., 2015). This study revealed that the production loss was the highest after one year of BVDV introduction, due to the increased number of abortions. Although a high number of abortions was also observed in the second year of our beef model (data not shown), the economic loss in our study was highest at year 3, when the increased number of abortions in year 1 and 2 resulted in a decreased number of fattening stocks able to be sold. Other factors, such as the culling age or price of fattening cattle, should also be noted when interpreting the direct losses of BVDV outbreak based on modelling studies.

The simulation models also provided valuable information on the anticipated benefit/cost ratios for different combinations of BVDV control options in dairy and beef farms. For the dairy farm, the annual testing of replacement breeding calves to find and eliminate PI animals showed the best benefit/cost ratio to control BVDV. Similar results have been reported in another study (Weir et al., 2014), and, indeed, testing and culling of PI animals as early as possible has been suggested as one of the most effective ways to rapidly eliminate BVDV within a farm (Houe et al., 2006). However, the testing and culling of PI breeding calves alone did not result in the highest benefits, which was likely because there was still significant risk of BVDV introductions onto the farm via contacts with infected stock on neighbouring farms or at off-site grazing locations for heifers. Therefore, implementing the annual testing of breeding calves with improving farm biosecurity through control measures, such as double fencing or movement restrictions, would be the most cost-efficient strategy for New Zealand dairy farms to control BVDV. Furthermore, improving farm biosecurity is likely to have positive benefits in reducing the between-farm transmission of other pathogens besides BVDV.

For the beef farm, the highest benefit/cost ratio was observed when a vaccination programme that involved administering only two doses before the first mating with annual boosting was implemented. However, this was due to the relatively low cost of control and it only resulted in a limited amount of benefit. Interestingly, additional control measures, such as annual testing or screening were not found to provide significant returns on investment for New Zealand beef farmers. This was because the annual testing or screening in this study was scheduled at the time of weaning. By the time PI calves were identified and culled, susceptible animals in the breeding herd had been exposed to those PI calves during the time window with the greatest risk of reproductive losses and generating PI calves. However, it is highly likely for most of beef farmers being reluctant to yard cows with calves at-foot prior to weaning due to the risk of injury to calves. Also, control measures, such as double fencing all farm boundaries, may not be logistically feasible for

all New Zealand beef farmers given the extensive farming conditions (Gates, Evans, et al., 2019). Therefore, we argue that vaccinating breeding stock for three doses before their first mating and maintaining immunity with annual boosters is likely to be a more economically beneficial and logistically feasible approach to control BVDV on New Zealand beef farms, even though it is still important to identify and remove PI animals from the herd and to prevent local transmission from neighbouring farms.

The model results also highlighted several strategies that were unlikely to be cost-effective for the typical New Zealand dairy and beef farms. A previous unpublished modelling study estimated the cost-benefit of different BVDV control options in New Zealand dairy farms and found that implementing almost any option would be economically beneficial for BVDV infected farms compared to doing nothing (Weir et al., 2014). The discrepancy between the previous and our current study was likely due to different model assumptions, such as the initial BVDV status of modelled animals (e.g. totally naïve or partially immune), type of animals that initially introduced BVDV (e.g. Trojan cows, PI cows, or calves), or conditions of BVDV introduction (e.g. endemic only, BVDV incursion only, or BVDV incursion under endemic condition). Therefore, one should consider the general context of BVDV introduction/infection when comparing the cost-benefit of different BVDV control options. Double fencing alone, for example, generated only a limited amount of benefit, however, we argue that this was due to the study design of forcing the PI animals to be initially present on the farm. Under circumstances where BVDV introduction via animal purchases was well-controlled and PI animals on neighbouring farms were the only source of transmission, double fencing alone could be one of the most cost-effective options to control BVDV (Reichel et al., 2008).

Our results showed that most BVDV control strategies involving annual screening tests were not economically beneficial for both dairy and beef farms. It suggests that a BVDV screening test based on BTM or blood samples from young heifers to detect either the virus or BVDV Ag would have poor performance to capture a recent BVDV outbreak on New

Zealand cattle farms. However, we evaluated the cost-efficiency of those strategies under the circumstance that BVDV outbreak occurred in a naïve farm due to the introduction of PI animals via Trojan cows. According to a recent study that analysed the laboratory submission data for testing BVDV in New Zealand, less than 12% of cattle farms showed the active circulation of the virus while more than 50% showed the evidence of recent exposure to BVDV (Gates, Han, et al., 2019). Therefore, the cost-efficiency of annual screening tests, especially at a national level, would likely be different given its low control cost and the majority of farms showing the lack of active circulation of BVDV with possibly high level of herd immunity. Currently, research work is in progress to estimate the benefit-cost ratios of different BVDV control strategies on a national scale.

There are several potential limitations in the model structure and assumptions that may have influenced the accuracy of our estimates. As with other modelling studies, our results were significantly affected by some model parameter values in the sensitivity analysis (see Section 4 of Appendix 3). For instance, several studies, including ours, have demonstrated that the mortality rate of PI animals had a significant impact on BVDV simulation results (Damman et al., 2015; Ezanno et al., 2007), and, interestingly, an observational report suggested that PI animals in New Zealand may survive longer than in other countries (Voges et al., 2006). Although we calibrated model parameters to be specific to the New Zealand cattle industries, we acknowledge that our results may not perfectly reflect the reality and rely heavily on model assumptions, such as the number of introduced Trojan cows, and within-herd BVDV transmission rate by a PI animal, as well as the values of some economic parameters, including reduced milk solid production in PI cows, and reduced live-weight gain per day in TI and PI animals. Nevertheless, we developed our model in such a way that any modelling settings (e.g. parameter values or management dates) can be flexibly modified for future applications. Another limitation was that we only considered a single BVDV control strategy being continually implemented during the 5-year simulation period. In reality, however, a farmer's biosecurity decisions can be

motivated by different factors, such as a farmer's knowledge about the disease, market-driven economic premiums or social and cultural significance (Garforth et al., 2013; Kristensen & Jakobsen, 2011; McAloon et al., 2017), which would affect their continual uptake of control strategies (e.g. shifting between different strategies or intermittent implementation of strategies). In future research, the efficiency of different control strategies could be investigated while considering other internal/external factors that would affect farmers' implementation of BVDV control measures.

5.6. Conclusion

Given that the economic impacts of BVDV outbreaks in naïve cattle herds in New Zealand have been estimated at NZ\$ 22.22 per MA per year on dairy farms and NZ\$ 41.19 per MA per year on beef farms, it is likely that the economic impacts of BVDV at the national level would be significant with the current high prevalence of BVDV infected cattle farms. The models suggest that BVDV control strategies involving the annual testing of breeding calves on dairy farms and the vaccination of heifers with annual boosters on beef farms are the most cost-effective control strategies at the farm-level although we acknowledge that farmers should still collaborate with veterinarians to determine the optimal strategy for their unique farm circumstances. Considering the importance of economic incentives for farmers to implement BVDV control, we believe that the findings of this study may facilitate the uptake of BVDV control measures among New Zealand cattle farmers by showing an economic benefit to BVDV control.

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Chapter 6

Modelling the cost-effectiveness of bovine viral diarrhoea virus elimination in New Zealand

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6.1. Abstract

Bovine viral diarrhoea virus (BVDV) is a widespread and economically important disease for most cattle producing countries worldwide. Several European countries have already implemented successful compulsory regional or national BVDV elimination programmes, which have led to complete elimination or significant reduction in disease prevalence. BVDV is currently under voluntary control in New Zealand and there is growing interest in evaluating whether national BVDV elimination would be cost-effective for the predominantly extensive pastoral production system. In this study, a stochastic individual-based metapopulation model was developed to simulate BVDV transmission within and between cattle herds and to explore the cost-effectiveness of different potential national BVDV elimination programmes. The model results suggest that BVDV transmission in New Zealand is maintained in an endemic state primarily through local spread between cattle herds grazed in adjacent pastures with direct farm-to-farm cattle movements contributing less to transmission events. The most cost-effective national BVDV elimination programmes involved (1) double fencing of shared boundaries for dairy farms accompanied by either annual testing and culling of persistently infected (PI) female

breeding calves or annual herd-level screening tests followed by a PI hunt, and (2) annual vaccination of breeding animals for beef breeding farms. The estimated benefit/cost ratio of the national BVDV elimination programmes ranged from 1.04 to 1.99 with the median time until reaching the break-even point ranging from 7 to 11 years depending on the programme, which highlights that BVDV elimination is both technically feasible and cost-effective for New Zealand pastoral production systems.

6.2. Introduction

Similar to many cattle producing countries worldwide, bovine viral diarrhoea virus (BVDV) is a widespread and economically important endemic disease in the New Zealand cattle industries (Anonymous, 2020). According to a recent analysis of BVDV diagnostic laboratory accession data, more than half of approximately 12,000 New Zealand dairy herds had high bulk tank milk (BTM) antibody levels indicating recent exposure to BVDV whilst approximately 7% also identified as RT-PCR positive implying active circulation of the virus within the milking herd (Gates, Han, et al., 2019). In the beef industry, several field studies have estimated that 40 to 60% of approximately 9,200 commercial farms have evidence of recent BVDV exposure with 22% likely containing persistently infected (PI) animals that actively shed BVDV throughout their lives (W. Cuttance & Cuttance, 2014; Gates, Han, et al., 2019; Han, Holter, et al., 2018). Given the high prevalence of exposed and actively infected farms, it is believed that BVDV causes significant economic impacts for the New Zealand cattle industries (Han, Weir, et al., 2018). Consequently, there has been growing interest in understanding whether implementing a national BVDV elimination programme could result in similar financial benefits to those observed in European countries with compulsory programmes.

The cornerstone of any national BVDV elimination programme is to remove PI animals from the cattle population through testing and culling in positive herds as well as enhancing

farm biosecurity to prevent the introduction/reintroduction of BVDV in negative herds (Lindberg et al., 2006). There are many different frameworks that can be used to achieve these goals based on the unique epidemiological, economic, and social circumstances present in each country. In Germany, for example, all calves must be tested for BVDV antigen (Ag) within 6 months of age to identify and remove any PI calves and farmers may also voluntarily choose to vaccinate their cattle against BVDV to protect against the formation of new PI calves (Moennig & Becher, 2018). Scotland, in contrast, uses annual herd-level screening of BVDV antibody (Ab) in serum samples of youngstock to identify herds with evidence of recent exposure and then performs a follow-up test of individual animals in positive herds to identify and remove PI animals. (Moennig & Becher, 2018). Regardless of the programme design, it is well-established that it is difficult to achieve meaningful reductions in BVDV prevalence without a high level of farmer compliance with disease control recommendations (Wernike et al., 2017). Therefore, a shift from the current *ad hoc* voluntary scheme to a compulsory systematic approach is necessary in order to eliminate BVDV in New Zealand.

It has been reported that the uptake of BVDV control measures in New Zealand is relatively low under the current voluntary scheme (Stewart, 2013). For example, a recent study suggested that only 1,600 beef farms in New Zealand are conducting annual screening tests for BVDV Ab (Gates, Han, et al., 2019). Although almost 75% of dairy farms are currently using breeding bulls that are either BVDV vaccinated or certified as BVDV free (Han, Holter, et al., 2018) and nearly 65% are currently conducting annual screening test of BVDV Ab using BTM (Gates, Han, et al., 2019), it is unknown what proportion of herds are taking additional measures to eliminate existing PI animals and prevent the formation of new PI calves. Possible barriers to control include farmers' poor awareness of BVDV, the logistical challenges of conducting sampling in beef herds, the low perceived impact of BVDV, and the high perceived cost of BVDV control (Gates, Evans, et al., 2019). Therefore, there is a recognised need to have a better estimation on the

cost-efficiency of national BVDV elimination in order to convince farmers and stakeholders of the value of BVDV elimination.

Disease simulation modelling is a common tool for exploring the transmission dynamics of BVDV in cattle populations (Damman et al., 2015; Viet et al., 2004) and for evaluating the cost-effectiveness of BVDV control programmes to provide decision-makers with guidance about the feasibility of BVDV elimination (Gethmann et al., 2019; Sekiguchi et al., 2018; Thulke et al., 2018). For instance, a recent simulation study suggested that a control programme based on Ag testing of all new born calves to identify and cull PI animals would reduce national BVDV prevalence to 0.01% with acceptable benefit/cost ratio in Germany (Gethmann et al., 2019). However, it is difficult to adopt the findings from models developed for one country to others with different farming systems because different demographic structures and contact patterns within cattle populations can significantly influence model results.

Although BVDV has been prevalent in New Zealand with its substantial economic impacts (Han, Weir, et al., 2018), there have been no published studies to simulate the transmission of BVDV or to evaluate the cost-effectiveness of national BVDV elimination programmes under a unique extensive pastoral cattle production system of New Zealand. With the aid of our recent studies about BVDV infection or transmission on New Zealand beef farms (Gates, Evans, et al., 2019; Han et al., 2019; Han, Holter, et al., 2018), we therefore aimed in this study (1) to develop a stochastic individual-based simulation model to explore the transmission dynamics of BVDV in the New Zealand cattle industries, and (2) to evaluate the cost-effectiveness of different approaches for national BVDV elimination.

6.3. Material and methods

To simulate the transmission dynamics of BVDV in the New Zealand cattle industries, we developed a stochastic individual-based metapopulation model with components to describe (1) the transition of individual animals between demographic and disease status within each farm (or subpopulation), (2) the spread of BVDV between farms through cattle movements and local spread from adjacent farms, and (3) the efficiency of different potential national BVDV elimination programmes. As there is currently no national database that provides complete information about the cattle population in New Zealand, we inferred the demographic structure of the New Zealand cattle industry (e.g. numbers of cattle farms and their location, herd size of each cattle farm, calving and mating dates of each cattle farm, and animal movements between those farms) by using available data in the National Animal Identification and Tracing (NAIT) system data. The model was developed using the C programming language, and the model was operated in a discrete time step of one week with a year consisted of 52 weeks (e.g. week 1: first week of January, week 52: last week of December).

6.3.1. Data processing

An extract of data from the NAIT system containing all known records from 01 January 2014 through 31 December 2016 was used to inform the demographic structure of cattle farms in New Zealand for the simulation model. The NAIT system has been in operation since 2012 with legislation requiring that all cattle owners or managers in New Zealand record the movements of individual cattle between different premises on the NAIT system within 48 hours of the movement occurring. Along with demographic details about the individual animals (e.g. animal ID, sex, date of birth), the owners or managers are also required to record other information about the movements including the premise ID, movement date, geo-coordinates of premises (as a point location), location type of the source and destination premises (e.g. farm, sale-yard, abattoir), and the farm type of the source and destination farms (e.g. dairy or beef farm). However, there are known issues

with the completeness and quality of these data due to the low level of compliance of New Zealand farmers (Edge & Kavalali, 2018; Jewell et al., 2016). Furthermore, there are currently no data fields to capture the total number of cattle on each farm, the size or boundary of farms, farm management information (e.g. planned start date of mating and calving), whether a farm is a commercial or hobby farm, the detailed production type of beef farms (e.g. beef breeding, beef fattening, bull breeding), and the purpose for the cattle movements (e.g. trade, show, bulls leased for breeding, heifer agistment).

To inform the demographic components of the simulation model, we therefore had to make inferences from the available data to (1) estimate the total number of active cattle farms in New Zealand, (2) allocate a herd size and planned start date of mating to each dairy and beef breeding farm, and (3) assign a point location to each farm to identify those that are likely adjacent to each other and therefore at risk of BVDV transmission through local spread. We also used the historical data on individual between-farm cattle movements to model BVDV spread in New Zealand through animal movements. Figure 6.1 provides a general overview of the data processing steps.

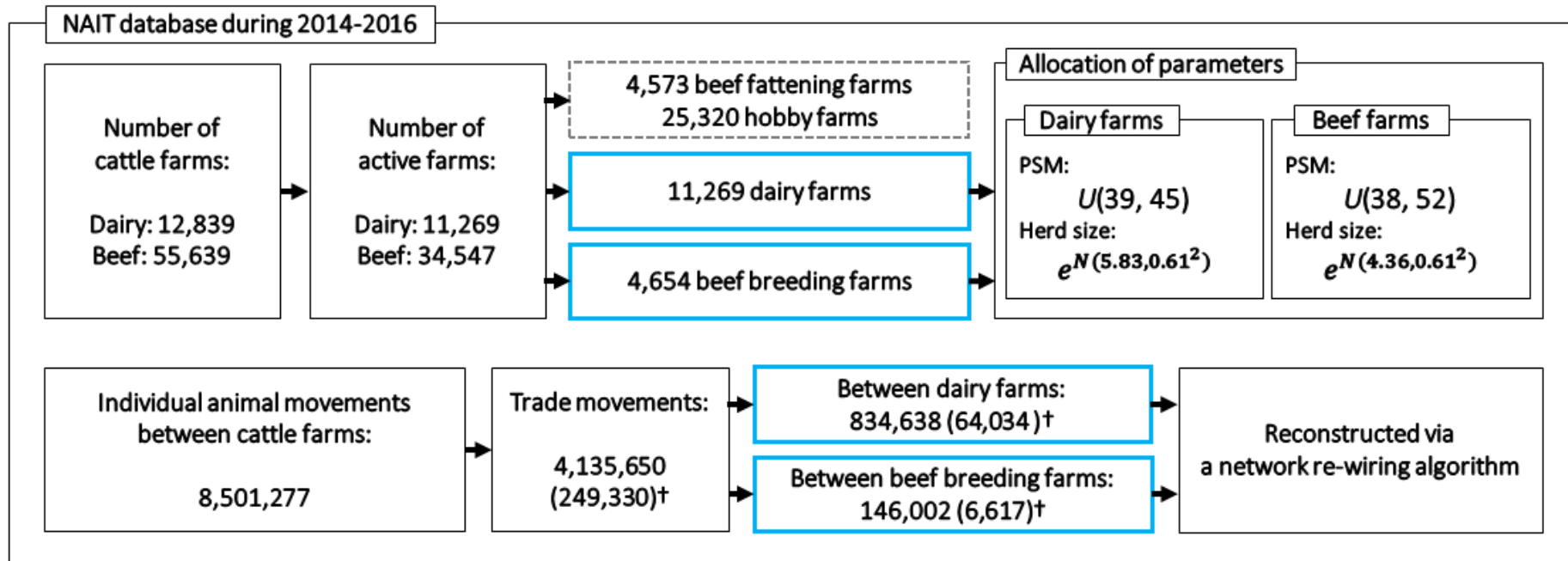


Figure 6. 1. Description of data processing. Data between 2014 and 2016 were extracted from the National Animal and Identification Tracing system in New Zealand. Only the data inside blue solid squares were explicitly incorporated into the national model, while the one inside a grey dashed square was indirectly modelled. † indicates the number of batches (i.e. a group of cattle moved from a source farm to a destination farm in a given week). Key: PSM, planned start of mating; U, Uniform distribution (lower, upper limit), N, normal distribution (mean, variance).

6.3.1.1. Number of cattle farms

The primary farms of interest for modelling the epidemiology and economics of BVDV are active commercial cattle farms with pregnant animals capable of generating new PI calves. Although farms that maintain only dry stock for grazing or fattening (including non-commercial or “hobby” farms) can also spread BVDV to neighbouring farms through fence-line contacts or other mechanisms of local transmission such as shared facilities and equipment (Niskanen & Lindberg, 2003), new PI animals cannot be generated on those farms, so the infections tend to be self-limiting. Therefore, the cattle farms that explicitly modelled in this study were dairy and beef breeding farms due to the presence of pregnant animals.

Some registered locations in the NAIT system only hold cattle temporarily and therefore have no current animals on the premises. However, the NAIT database does not contain a specific field to indicate whether a location is actively farming or grazing cattle. We therefore made the simplifying assumption that cattle farms were active if at least one animal was sent from the location to an abattoir from 01 January 2014 to 31 December 2016. This resulted in an estimate of 11,269 active dairy farms and 34,547 active beef farms. All dairy farms were assumed to be commercial since the estimate was close to the number of dairy farms in New Zealand reported by other sources (Jewell et al., 2016; Van Andel et al., 2018).

Currently, it is believed that there are approximately 9,200 commercial beef farms, with nearly half of them being beef breeding farms and the rest being beef fattening farms (*pers. comm.* Rob Davison). We therefore identified the commercial beef farms from 34,547 active beef farms on the NAIT database based on the number of cattle slaughtered (i.e. sent to abattoirs). We first assumed that commercial beef farms were those with (1) at least one animal slaughtered per year, and (2) more than x number of animals slaughtered between 2014 and 2016. We then adjusted the value of x until the number of identified commercial beef farms was approximately 9,200. Following these assumptions, 9,227 farms (26.7%)

were categorised as commercial beef farms and they each sent > 75 animals to abattoirs during 2014-2016. The remaining beef farms (25,320; 73.3%) were categorised as hobby farms. For the rest of this article, dairy and beef farms indicate commercial dairy and beef farms, respectively. To further categorise beef farms into breeding or fattening farms, we explored the patterns of movements onto these farms during 2014-2016. In the New Zealand cattle industries, it is common for beef fattening farms to purchase male dairy calves and surplus female dairy calves to raise for slaughter, whereas it would be relatively uncommon for beef breeding operations to purchase these cattle. Therefore, we categorised the farm as a beef fattening farm if any animals from dairy farms were moved to the beef farm during 2014-2016. Based on this assumption, 4,654 beef farms (50.4%) were categorised as beef breeding farms and 4,573 farms (49.6%) were categorised as beef fattening farms. The number of categorised beef breeding farms was well-matched to industry estimates of the number of commercial beef breeding farms.

Rather than explicitly modelling the animals on beef fattening and hobby farms, the impact of those farms on national BVDV transmission was indirectly incorporated into the model by assuming that all fattening and dry stock on those farms contributed to the background prevalence of PI animals in New Zealand, and therefore, contributed to the risk of BVDV transmission through local spread on dairy or beef breeding farms or at off-site grazing locations (i.e. distant premises where dairy replacement heifers are sent for grazing and breeding). Details about how animals on beef fattening or hobby farms were modelled is provided in Appendix 4 (see Section 1).

6.3.1.2. Herd size

To allocate a realistic herd size to each dairy and beef breeding farm in the study, we assumed that the distribution of number of mixed-age cows (MA) in New Zealand dairy and beef farms followed the exponential distribution of a random variable that normally distributed (i.e. $e^{N(\mu_D, \sigma^2)}$ and $e^{N(\mu_B, \sigma^2)}$, respectively), where $N(\cdot, \cdot)$ was the normal distribution of random variable, μ_D and μ_B were the mean of the normal distribution for

dairy and beef farms, respectively, and σ was the standard deviation of the normal distribution. We inferred the value of μ_D , μ_B , and σ as 5.83, 5.02, and 0.61, respectively, by fitting the simulated MA herd sizes to the reported distribution of MA herd size in the dairy industry (Anonymous, 2018a) and to the reported total number of MAs in the beef industry (Anonymous, 2018c) using an approximate Bayesian computation-sequential Monte Carlo method (Ellen Brooks-Pollock et al., 2014; Toni et al., 2009). Using the inferred distributions, we generated 11,269 random MA herd size numbers for the dairy farms and 4,654 random MA herd size numbers for the beef farms, and then sorted each list in descending order from largest to smallest. Next, we ranked the dairy and beef farms from largest to smallest based on the number of cattle sent to abattoirs from 2014 to 2016, and then matched them to the MA herd size with the same rank (e.g. assigning the largest MA herd size to the farm with the largest number of animals sent to abattoirs). This was based on the assumption that the number of animals sent to abattoirs from each farm was a proxy for the total cattle population on the farm. We further assumed that the allocated MA herd size for each dairy and beef breeding farm was constant during the simulation period. Detailed information about the process for inferring herd size distributions is provided in Appendix 4 (see Section 2).

6.3.1.3. Planned start of mating

To allocate the planned start of mating (PSM) to dairy and beef breeding farms, we first generated the distribution of week of conception for all New Zealand dairy and beef cattle by (1) extracting the birth week of all individual dairy and beef cattle, respectively, via available NAIT data, and (2) calculating back 40 weeks from the reported birth week. Given the mostly seasonal nature of the New Zealand cattle farming industries, approximately 95% of the conceptions occurred from late September to early January (week 39 ~ week 2) for dairy cattle and from mid-September to late February (week 38 ~ week 9) for beef cattle. We assumed conception mostly occurred during the first 9 weeks of the mating period for both dairy and beef cows, so the PSM should start at least 9 weeks

before the end of the period for most conceptions to occur. Based on this assumption, the PSM was randomly chosen between week 39 and week 45 (reversing 9 weeks from week 2) for a dairy farm and between week 38 and week 52 (reversing 9 weeks from week 9) for a beef breeding farm.

6.3.1.4. Adjacent farms

To identify farms that were at risk of BVDV transmission through local spread, we needed to calculate the Euclidean distance between farm locations. Among the active cattle farms, 1,452 (12.9%) of 11,269 dairy farms, 572 (12.3%) of 4,654 beef breeding farms, 609 (13.3%) of 4,573 beef fattening farms, and 2,785 (11.0%) of 25,320 hobby farms did not have a recorded point location in the NAIT system. Rather than exclude them from the analysis, we randomly allocated coordinates for these farms based on the overall spatial distribution of cattle farms with respective production type in New Zealand. The spatial distribution of dairy and beef breeding farms in New Zealand is illustrated in Figure 6.2. Once all cattle farms had a unique point location, we defined two farms as being adjacent if the Euclidean distance between point locations was ≤ 3.0 km, where the distance of 3.0 km was calibrated to meet the simulated prevalence of BVDV positive farms to the reported prevalence (see Section 2 in Appendix 4).

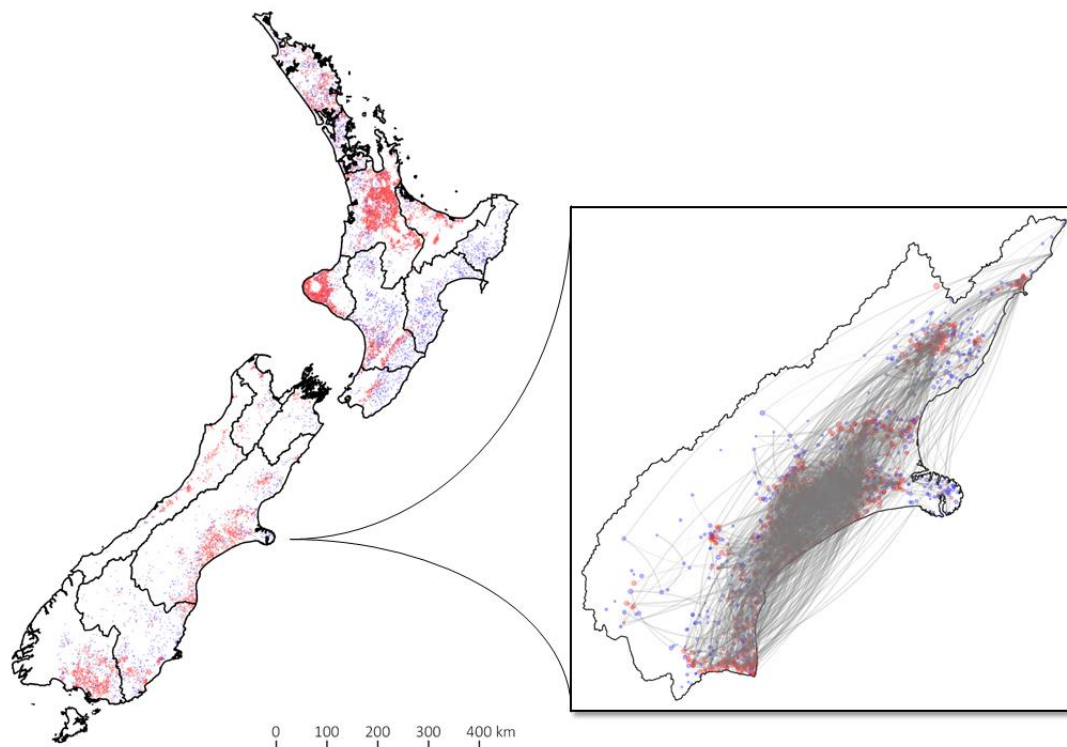


Figure 6. 2. Distribution of cattle farms in New Zealand (left) and the Canterbury region (right). Left & Right: Red and blue dots indicate dairy and beef farms, respectively, with the size of dot was proportional to the mixed-age cow herd size for each farm. Right: Grey lines in the Canterbury region indicate aggregated animal movements via trade during 2014-2016 (direction was ignored for the simplicity of illustration).

6.3.1.5. Cattle movements

Given that cattle movements for temporary grazing are often unrecorded in the NAIT system (Edge & Kavalali, 2018), we focused only on the movement of cattle for trade purposes. However, the impact of off-site grazing movements on BVDV transmission was implicitly incorporated into the model (see “6.3.2.3. BVDV transmission within a farm” and “6.3.2.4. BVDV transmission between farms” sections). Winter grazing (i.e. when dried-off, pregnant dairy cows are temporarily grazed on a distant location for several weeks during winter) was ignored because any pregnant cattle would be outside the gestational risk period for generating PI animals even if they were exposed to BVDV at the distant location.

To identify the trade movements, we first extracted all individual cattle movements between any cattle farms where the source and destination farms had a different premise ID. If an individual animal returned to the source farm in a subsequent movement, both movements were classified as non-trade movements and discarded. Next, we aggregated the individual cattle movements into weekly batch level movements (i.e. a group of cattle moved from a source farm to a destination farm in a given week), and further discarded any batches where the total number of cattle moved from a source farm to a destination farm during May and June was > 50 . This was done to avoid including any possible migration movements of “*sharemilkers*” (farmers who own only the cattle and lease land for grazing, therefore they move between farm-premises periodically – generally around June 1st). Finally, we extracted only the movements between dairy farms and beef breeding farms since these movements potentially involved the movement of breeding animals. We ignored any movements between farms where pregnant animals cannot exist.

According to the NAIT database during 2014-2016, there were 834,638 individual cattle movements (in 64,034 batches) between dairy farms, 146,002 individual cattle movements (in 6,617 batches) between beef breeding farms, and 28,509 individual cattle movements (in 2,517 batches) from beef breeding to dairy farms. From these movements, we discarded those that occurred from beef breeding to dairy farms since these were most likely beef breeding bulls that were hired only for the mating period, and a high proportion of breeding bulls are BVDV tested and vaccinated (Han, Holter, et al., 2018). We then used the information about the number of farms that moved cattle and the number of cattle moved between farms with the same production type over the 3 years (156 weeks in total) for the network re-wiring algorithm to reconstruct the cattle movements in the national model (see Section 2 in Appendix 4).

6.3.2. National model

6.3.2.1. Demographics of cattle within a farm

Most cattle farms in New Zealand are seasonal calving herds. The planned start of calving (PSC) is generally between mid-July and early August for dairy farms and between August and November for beef breeding farms. Calving season lasts for approximately 9 to 12 weeks for dairy farms, while ≤ 9 weeks is common for beef farms. Off-site grazing is relatively common practice only for dairy farms.

We used the within-farm models that were developed in Chapter 5 with some modified parameter values due to the difference in time units (day vs. week). For a dairy farm, all cattle were divided into 4 management groups in the within-farm model: calves (C: from birth to until 1 week old), young replacement heifer (YH: from 1 week old to off-site grazing), breeding replacement heifers (BH: heifers at the off-site grazing location), and mixed-aged cows (MA). We assumed that a number of female calves that equivalent to approximately 20.0% of allocated MA herd size was retained as YHs, and the rest of female calves and all male calves were either culled (i.e. bobby calves) or sold to beef fattening farms or hobby farms (see Section 1 in Appendix 4). The duration of off-site grazing was assumed to be 52 weeks, and the week for YHs (or BHs) to be sent for (or returned from) off-site grazing was randomly chosen between weeks 18 and week 22 (between late April and late May) for each dairy farm. For simplicity of modelling, we also assumed that the dry period for MAs on each dairy farm started at the same week when BHs were returned from off-site grazing. For the within-farm model of beef breeding farms, we divided the beef population into 7 management groups: fattening heifers (FH) and steers (FS) for fattening, young replacement heifers (YH: heifers in this herd are bred for the first time), breeding replacement heifers (BH: heifers in this herd calve for the first time and are bred for the second time) and their calves (CH), and mixed-age cows (MA) and their calves (CM). We assumed that most of the transitions between beef management groups occurred on the week of weaning (31 weeks after the PSC), however, all BHs joined the MA herd

three weeks before the week of weaning and pregnancy scanning was performed one week before the week of weaning. Breeding bulls were not considered in the simulation models for both dairy and beef farms since most of them are tested for, and vaccinated against, BVDV (Han, Holter, et al., 2018). Parameter values for the national model are provided in Table 6.1. More detailed information about the demographic structure and events of within-farm models is provided in Appendix 4 (see Section 2) and Chapter 5. Description of demographic events in both dairy and beef farms in the national model is provided in Figure 6.3.

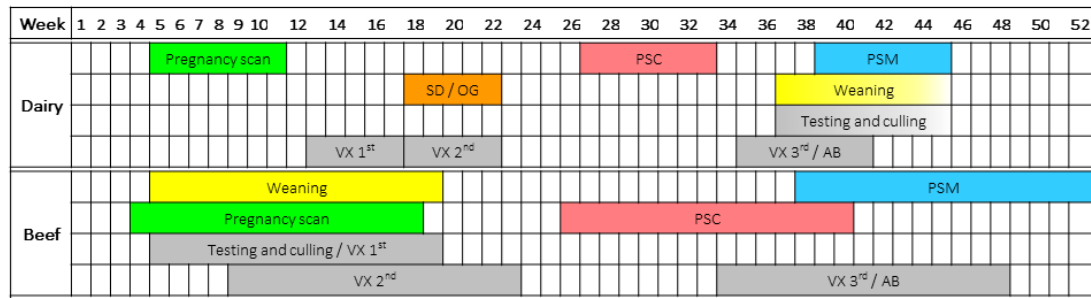


Figure 6. 3. Illustration of demographic events in the New Zealand dairy and beef industries. SD, start of dry-off period; OG, week to move from/to off-site grazing location; PSC, planned start of calving; PSM, planned start of mating; Testing and culling, annual testing and culling of female breeding calves; Screening test, annual screening of bulk tank milk (dairy) or serum samples of 15 young heifers (beef) followed by a PI hunt; VX 1st, 2nd, and 3rd, the first, second, and third BVDV vaccination, respectively; AB, annual boosting of BVDV vaccination.

Table 6. 1. Description of parameter values used in the national BVDV simulation model.

Parameters	Value (unit)	Reference
Demographic component		
Natural abortion	0.001 (/week)	Hickson et al., (2012; Weston et al., (2012)
Natural mortality by weaning;		E. Cuttance et al., (2017); Hickson et al., (2012); C. A. Morris et al., (1986)
Dairy calves	0.010 (/week)	
Heifer-born beef calves	0.006 (/week)	
Cow-born beef calves	0.002 (/week)	
Start of puberty	54 (weeks old)	Hickson et al., (2012); McNaughton et al., (2002)
Length of oestrus cycle	3 (weeks)	Hickson et al., (2012); Olds & Seath, (1953)
Probability of oestrus detection (dairy only)	0.786	Anonymous, (2018a)
Probability of conception;		Anonymous, (2018a); Geenty & Morris, (2017)
By artificial insemination (dairy only)	0.72	
By natural mating with bulls	D: 0.50 / B: 0.72	
Disease component		
Transmission rates of;		Han et al., (2019); Smith et al., (2014); Viet et al., (2004)
Within-herd by a PI animal (β_P)	D: 0.50 / B: 0.11	
Within-herd by a TI animal (β_T)	$\beta_P \times 0.05$	
Between-herd by a PI animal (β_B)	$\beta_P \times [0.10, 0.20, 0.40]$ *	
Between-farm by a PI animal (β_F)	$\beta_P \times [0.10, 0.20, 0.40]$ *	
Euclidean distance for adjacent farms (d)	[2.0, 3.0, 4.0] *(km)	
Duration of passive immunity via maternal antibody	$N(22, 4)$	Han et al., (2019)
Duration of recovery of TI animals	$U(1, 3)$	Liebler-Tenorio et al., (2004); Müller-Doblies et al., (2004); T. Sandvik et al., (1997)
Mortality of PI animals	0.013 (/week)	Ezanno et al., (2007)
Reduced conception of TI and PI animals	0.68	McGowan et al., (1993); Whitmore et al., (1981)

Probability of abortion if;		Grooms et al., (2007); McClurkin et al., (1984); McGowan et al., (1993); Viet et al., (2004); Walz et al., (2017, 2018)
Dams were infected during 0 ~ 41 days of gestation	0.121	
Dams were infected during 42 ~ 150 days of gestation	0.174	
Dams were PI and at 0 ~ 41 days of gestation	0.060 (/week)	
Dams were PI and at 42 ~150 days of gestation	0.090 (/week)	
Probability of BVDV infection status of calves if dams were infected during 42 ~ 150 days of gestation;		Viet et al., (2004)
PI calves	0.934	
Passively immunized calves via maternal antibody	0.033	
Recovered calves	0.033	
<hr/>		
Control component		
Sensitivity of BVDV RT-PCR	0.999	Lanyon et al., (2013)
Specificity of BVDV RT-PCR	0.999	
Sensitivity of BVDV ELISA antigen test	0.835	
Specificity of BVDV ELISA antigen test	0.994	
Probability of vaccination outcome (after second dose);		Chapter 5
Vaccine failure	0.193	
Insufficient fetal protection (V_I)	0.734	
Sufficient fetal protection (V_S)	0.072	
Probability of abortion if;		Chapter 5
V_I dams were infected during 0 ~ 41 days of gestation	0.773	
V_I dams were infected during 42 ~ 150 days of gestation	0.762	
Probability of BVDV infection status of calves if V_I dams were infected during 42 ~ 150 days of gestation;		Chapter 5
PI calves	0.232	

Passively immunised calves via maternal antibody 0.559

Recovered calves 0.209

* Values in square bracket were tested for the parameter calibration. Bold italic values were the calibrated parameter values that used for the national model.

Key: *PI*, Persistently infected; *TI*, Transiently infected; *N*, Normal distribution; *U*, Uniform distribution; *D*, Dairy; *B*, Beef.

6.3.2.2. Movement reconstruction

When we incorporated the real movement data directly into the model to add or remove individual cattle from a herd, the size of the cattle populations on some farms often either increased or decreased substantially over time. To address this problem, we used a network re-wiring algorithm that connected pairs of farms that moved cattle in any given week based on the procedure described below.

For a given week, the same number of dairy (or beef breeding) source and destination farms were selected in a random order. The number of selected farms was chosen as the larger number between the numbers of dairy (or beef breeding) source and destination farms in the relevant week. For the first source farm, the probability of trading with each destination farm was calculated based on a multivariable logistic regression model that considered the distance, herd size, and tuberculosis movement restriction status (see Section 2 in Appendix 4); we then assumed that animals were sold to the destination farm with the highest probability. Once those two farms were connected, the algorithm moved on to the next source farm, calculated the probability of trading with each of the remaining destination farms, and connected it to the farm with the highest probability. The re-wiring process for a given week was finished once there were no remaining farms to either sell or purchase cattle.

Once the farms trading cattle were paired in each week, animals from a source farm were randomly selected and moved to a destination farm regardless of their BVDV infection status. However, we assumed that calves from both dairy and beef breeding farms and BHs from dairy farms were not eligible for trade due to their age and off-site grazing, respectively. The probability of an individual animal being selected for trade from a source farm was assumed to be the same for every animal (except those that were not eligible) from all the source farms in a given week, and varied between weeks to match the total number of animals traded in a given week in the simulation to the number of cattle traded per week in a relevant week from NAIT data.

6.3.2.3. BVDV transmission within a farm

To simulate BVDV transmission on each farm, we categorised the BVDV infection status of individual animals into the following five mutually exclusive disease status: passively immunized via ingestion of maternal antibodies through colostrum (M), susceptible (S), transiently infected (TI), persistently infected (PI), or recovered (R). In the simulation, the probability of BVDV infection for each susceptible animal was determined by the numbers of PI and TI animals within the same management group, PI animals in different groups on the same farm, PI animals on adjacent farms, and the background prevalence of PI animals. The probability of a susceptible individual in a group i being infected on a given week (P_{inf}) was;

$$P_{inf} = 1 - e^{-(\lambda_i^w + \lambda_i^b)}$$

$$\lambda_i^w = \beta_P \frac{PI_i}{N_i} + \beta_T \frac{TI_i}{N_i} + \beta_B \sum_j^n \frac{PI_j}{N_i N_j}, (i \neq j)$$

where λ_i^w was the force of within-farm BVDV infection for an animal in group i , β_P and β_T were within-group transmission rates by a PI and TI animal, respectively, β_B was a between-group transmission rate by a PI animal, N_i , PI_i , and TI_i were the group size, the number of PI and TI animals in group i on a given week, respectively, and n was the number of groups in a farm. The value of β_P for dairy and beef farms was 0.50 and 0.11, respectively (Han et al., 2019; Viet et al., 2004), and we assumed that β_T was 0.05 times of β_P . We calibrated the value of β_B as 0.40 times of β_P based on a sub-model that simulated BVDV transmission within cattle farms in the Canterbury region (see Section 3 in Appendix 4). Since animals in the BH herd on dairy farms were grazed at an off-site location, we assumed that any between-herd transmission to and from these animals could not occur. Parameter values for BVDV transmission are provided in Table 6.1, and detailed information about BVDV transmission is provided in Chapter 5.

6.3.2.4. BVDV transmission between farms

The contribution of cattle movements to between-farm BVDV transmission was explicitly modelled by moving individual animals to another farm and updating the force of infection on the destination farm at the following week according to the demographic group and BVDV infection status of the animals that were moved. For local spread, we assumed that any PI animals on adjacent dairy or beef breeding farms affect BVDV infection of individuals on a given farm. The impact of BVDV transmission of beef fattening or hobby farms to local spread was also incorporated by accounting for the total number of those farms that adjacent to a given farm. For an individual animal in group i on a given farm, the force of between-farm BVDV infection (λ_i^b) was calculated as:

$$\lambda_i^b = \beta_F \sum_k^l \frac{\mathbb{P}_k}{N_i \mathbb{N}_k} + \beta_F \frac{r \rho_B}{N_i}$$

where β_F was a between-farm transmission rate by a PI animal, N_i was the number of animals in group i on a given week, l was the number of adjacent dairy and beef breeding farms, \mathbb{N}_k and \mathbb{P}_k were the total number of cattle and PI animals on an adjacent dairy or beef breeding farm k on a given week, respectively, r was the total number of adjacent beef fattening and hobby farms, and ρ_B was the background prevalence of PI animals in New Zealand on a given week. Off-site grazed dairy BH herds were assumed to be exposed to ρ_B only (with $r = 1$), and we calibrated the value of β_F as 0.20 times of β_P based on the result of sub-model for the Canterbury region. Details about the estimation of β_F and ρ_B is provided in Appendix 4 (see Section 3).

6.3.2.5. Model initiation

We started the simulations of the national model by generating an initial population for each farm. For a dairy farm, x number of pregnant MAs, y number of pregnant BHs, and y number of YHs were generated at the beginning of the simulation where x and y were 100.0% and 20.0% of allocated MA herd size, respectively. For a beef breeding farm, we initialised

the model with x number of pregnant MAs, y number of pregnant BHs, y number of YHs, and z number of both CMs and CHs where x , y and z were 100.0%, 20.0% and 91.0% of the allocated MA herd size, respectively. Those proportions of initial animals on each farm was determined based on the national statistics for cattle populations in New Zealand (Anonymous, 2018c).

Simulations started on week 1 of year 1, and the model was run for 5 years as a demographic burn-in period to stabilise the population demography. We then introduced BVDV to 796 randomly selected farms (5% of sum of dairy and beef breeding farms) by converting one new born calf to a PI animal (dam of the calf was then converted to recovered) for each farm. The model was then run for another 40 years (from year 6 to year 45) to allow BVDV to establish an equilibrium state across New Zealand reflective of the current endemic situation on New Zealand cattle farms. Outputs from the last 15 years of simulation were visually inspected to confirm the stability of the population demographics. We also confirmed that the observed prevalence of BVDV positive farms (i.e. proportion of cattle farms carrying at least one PI animal) matched estimates of the national disease prevalence during the last 15 years (Gates, Han, et al., 2019). The overall modelling process is illustrated as a flow chart in Appendix 4 (see Section 4).

6.3.3. National BVDV elimination programmes

In this study, we examined eight different national BVDV elimination programmes (Table 6.2). Since we confirmed the equilibrium state of BVDV between year 31 and year 45, we implemented each programme continually from week 1 of year 31 onward and the effect of the programme was monitored over the following 15 years which encompassed the timeframe reported for the elimination of BVDV in European countries (Ståhl & Alenius, 2012). Each possible programme was a combination of five different potential control options: (1) movement restriction, (2) annual testing of female calves selected for

replacement heifers and culling of identified PI animals, (3) annual screening test using bulk tank milk (for dairy farms) or serum samples of 15 YHs (for beef breeding farms), followed by a PI hunt (i.e. identifying and culling of possible PI animals) of all breeding female animals if the screening test was positive, (4) vaccination of breeding animals with annual boosting, and (5) double fencing (Table 6.2).

6.3.3.1. Movement restriction

For the movement restriction option, female breeding animals were tested using a BVDV Ag ELISA before departing the source farm, and only the negative animals were moved to destination farms. We also restricted the movement of off-site grazing (both sending and returning) by conducting BVDV Ag ELISA testing four weeks before the week of off-site grazing and only allowing the movement of test negative heifers.

6.3.3.2. Annual testing and culling

The same Ag ELISA was used for the annual testing of female breeding calves, and the test was conducted at 1 week old for dairy calves and at weaning for beef calves. When animals were tested using the BVDV Ag ELISA (in both movement restriction and annual test and cull of breeding calves), test positive animals were re-tested after 4 weeks (i.e. follow-up test) using the same ELISA, and culled immediately if the result of the follow-up test was positive. When the follow-up test was conducted, dams of the positive animals were also tested using the BVDV Ag ELISA, and culled immediately if the result was positive. Sensitivity and specificity of the BVDV Ag ELISA was assumed to be 83.5% and 99.4%, respectively (Lanyon et al., 2013).

6.3.3.3. Annual screening test

Given the endemic status of BVDV, the annual screening test was aimed at detecting BVDV using RT-PCR on bulk tank milk (BTM) for dairy farms or BVDV Ag using the ELISA on individual serum samples from 15 randomly chosen YHs for beef farms. If the RT-PCR test of BTM (dairy farms) or the Ag ELISA of at least one serum sample (beef

breeding farms) was positive, a PI hunt was conducted after 4 weeks (dairy farms) or 3 weeks (beef breeding farms) from the initial screening test. The screening test was conducted at the PSM for dairy farms and 3 weeks before the week of weaning for beef breeding farms. We assumed the sensitivity and specificity of RT-PCR to detect BVDV on BTM was perfect (Table 6.2) (Lanyon et al., 2013).

6.3.3.4. Vaccination

Given that only killed BVDV vaccines are available in New Zealand, we assumed that the length of fetal protection after the second and third dose (and thereafter) of BVDV vaccination were 26 and 52 weeks, respectively, and the efficacy of vaccination differed according to the number of doses that had been given (i.e. no effects after the first dose, and partial protective effect after the second dose and thereafter). We scheduled BVDV vaccination to confer 52 weeks of immunity to female breeding animals before their first mating. On dairy farms, all YHs were vaccinated for the first time 4 weeks before being sent for off-site grazing, the second time at the week they were sent to off-site grazing, and the third time 4 weeks before the PSM. On beef breeding farms, the first BVDV vaccination was administered to calves at weaning, the second vaccination 4 weeks later, and the third time 4 weeks before the PSM. Annual boosters were administered to MAs (and BHs in case of beef farms) at 4 weeks before the PSM for both dairy and beef farms. Parameter values related to BVDV vaccination are provided in Table 6.1, and detailed information about the parameterisation of BVDV vaccination is provided in Chapter 5

6.3.3.5. Double-fencing

We assumed that the double fencing option prevented any local spread of BVDV from PI animals on adjacent farms as well as from the exposure to background prevalence of PI animals.

Table 6. 2. Description of the national BVDV elimination programmes.

Programme	Dairy farms	Beef breeding farms
P1	Movement restriction	Movement restriction
P2	Annual testing and culling	Annual testing and culling
P3	Annual screening test	Annual screening test
P4	Double fencing	Double fencing
P5	Vaccination with annual boosting	Vaccination with annual boosting
P6	Double fencing	Vaccination with annual boosting
P7	Annual testing and culling + Double fencing	Vaccination with annual boosting
P8	Annual screening test + Double fencing	Vaccination with annual boosting

6.3.4. Model Outputs

In order to compare the effect of implementing different national BVDV elimination programmes, we first monitored the number of dairy and beef breeding farms that were positive to BVDV during the simulation period (from week 1 year 31 to week 52 year 45) by assuming that a farm was BVDV positive if there was at least one PI animal on the farm over that time period. We then measured the effectiveness of each national BVDV elimination programme as the % reduction in the number of BVDV positive farms at the end of the simulation period compared to that one at the beginning of the simulation period. The number of BVDV positive farms was also divided by the total number of dairy and beef breeding farms (i.e. 15,923) to estimate the national prevalence of BVDV positive farms, and we recorded the last timepoint that the national prevalence of BVDV positive farms was 2.0% in order to evaluate how quickly the elimination programmes reduced the burden of the disease (i.e. time-efficient).

Next, we examined the gross farm revenue (GFR) and BVDV control cost per year for all dairy and beef breeding farms to conduct cost-benefit analysis of BVDV elimination programmes for the New Zealand cattle industries. We calculated the annual benefit as the sum of (1) difference between GFRs with and without various BVDV elimination programmes per year, and (2) an annual control expenditure which was the total cost currently spent for voluntary BVDV testing by individual cattle farmers (Yarnall &

Thrusfield, 2017). The annual control expenditure was estimated as NZ\$ 12.7 million per year based on the total number of BVDV-related laboratory tests during 2016-2017 (Gates, Han, et al., 2019). The parameter values to estimate the annual GFR and BVDV control cost for New Zealand dairy and beef breeding farms are provided in Chapter 5. For each programme, we accumulated the annual benefit and BVDV control cost with a discounting rate of 1.9%, and examined (1) whether or when a BVDV elimination programme had a break-even point (i.e. a timepoint that the accumulated benefit was the same as or larger than the accumulated BVDV control cost), and (2) the benefit/cost ratio of BVDV elimination programmes. The benefit/cost ratio was calculated as the accumulated benefit divided by the accumulated BVDV control cost in the final year. Due to the high computational burden of running a national model, each elimination programme was iterated only 100 times, however, we confirmed that the 100 iterations adequately captured the variation caused by the stochasticity (see Section 5 in Appendix 4).

6.4. Results

6.4.1. Effectiveness of national BVDV elimination programmes

The effectiveness of different national BVDV elimination programmes is provided in Table 6.3, and the national prevalence of BVDV positive farms over the simulation period is illustrated in Figure 6.4, 6.5, and 6.6. Implementing the BVDV elimination programmes of movement restriction, annual testing and culling, and annual screening test for both dairy and beef farms (P1, P2, and P3, respectively) was not effective as sole control strategies to reduce the national prevalence of BVDV infections. Any programme that involved double fencing almost eliminated BVDV from the New Zealand dairy industry (Table 6.3). For the beef breeding farms, any elimination programme that included vaccination was generally effective in reducing BVDV infections, with the lowest number of BVDV positive farms being observed when P7 was implemented (Table 6.3).

The last week that the national prevalence of BVDV positive farms being 2.0% with respect to different national BVDV elimination programmes is provided in Table 6.3. The most time-efficient BVDV elimination programmes were P7 and P8 with the median of 7 years (364 and 365 weeks, respectively) to reach national BVDV prevalence of 2.0%, and this was followed by P6 (median of 7.2 years or 372 weeks) and P5 (median of 8.9 years or 463 weeks).

Table 6. 3. Median (95% prediction interval) of the % reduction in the number of BVDV positive farms between the beginning and end of simulation period (i.e. % reduction) and the last week that the national prevalence of BVDV positive farms being 2.0% (i.e. last week for 2.0%) with respect to implementing different national BVDV elimination programmes.

Programme	% reduction		Last week for 2.0%
	Dairy (%)	Beef breeding (%)	
P1	38.0 (37.8, 38.3)	4.6 (4.4, 4.7)	-
P2	71.7 (69.5, 73.3)	1.8 (1.3, 2.4)	-
P3	20.7 (20.4, 21.0)	5.7 (4.7, 5.8)	-
P4	99.8 (99.4, 100.0)	69.7 (68.5, 70.6)	-
P5	89.5 (88.6, 91.2)	89.6 (88.8, 90.4)	563 (463, 772)
P6	99.8 (99.3, 100.0)	89.8 (88.6, 93.7)	411 (372, 419)
P7	100.0 (100.0, 100.0)	98.7 (94.3, 99.4)	364 (372, 414)
P8	100.0 (99.9, 100.0)	93.6 (89.3, 98.4)	365 (375, 414)

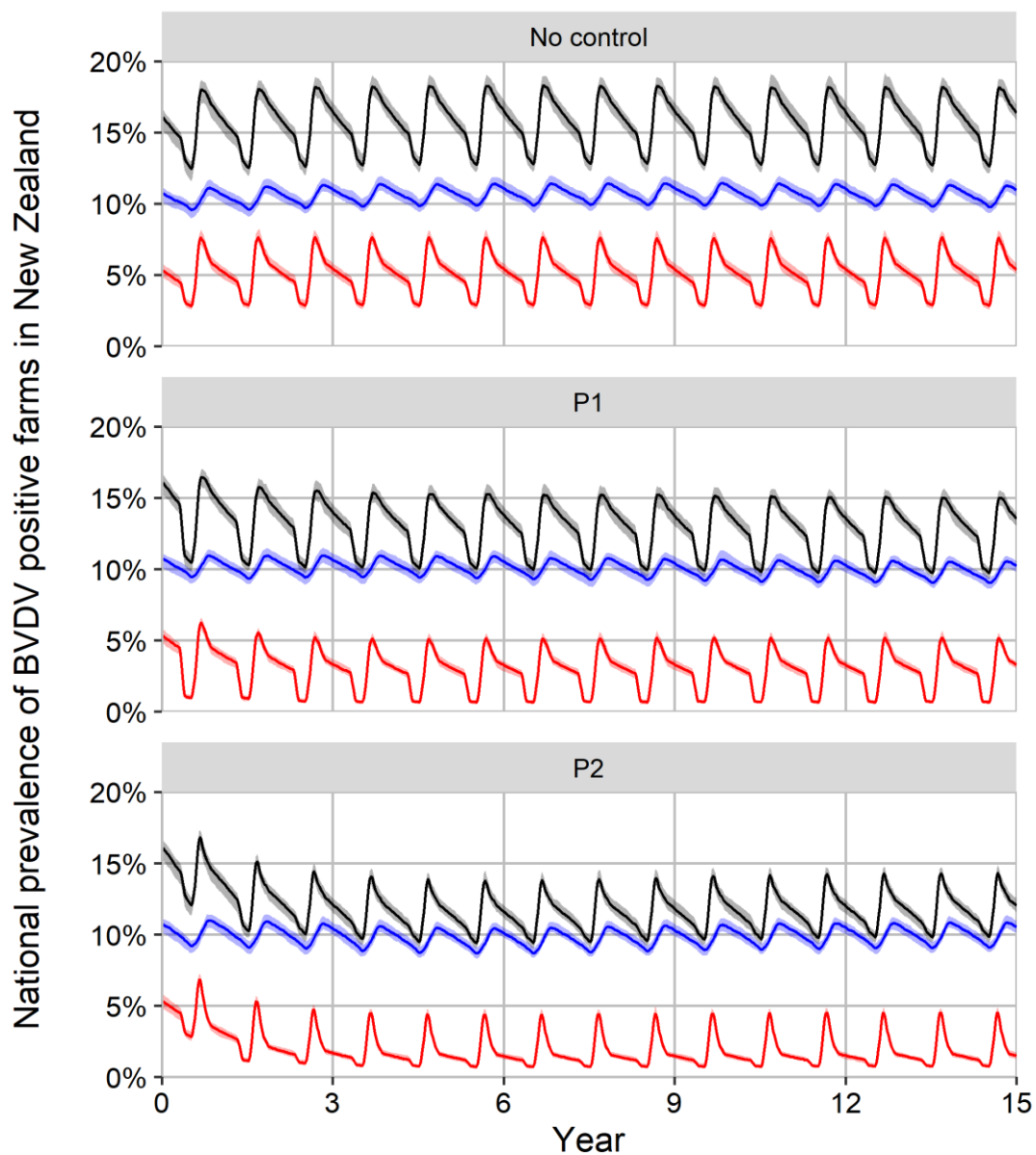


Figure 6. 4. National prevalence of BVDV positive farms (black) with respect to implementing no BVDV elimination programme, movement restrictions (P1), and annual testing and culling of breeding female calves (P2) for both dairy and beef breeding farms. Red and blue lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively.

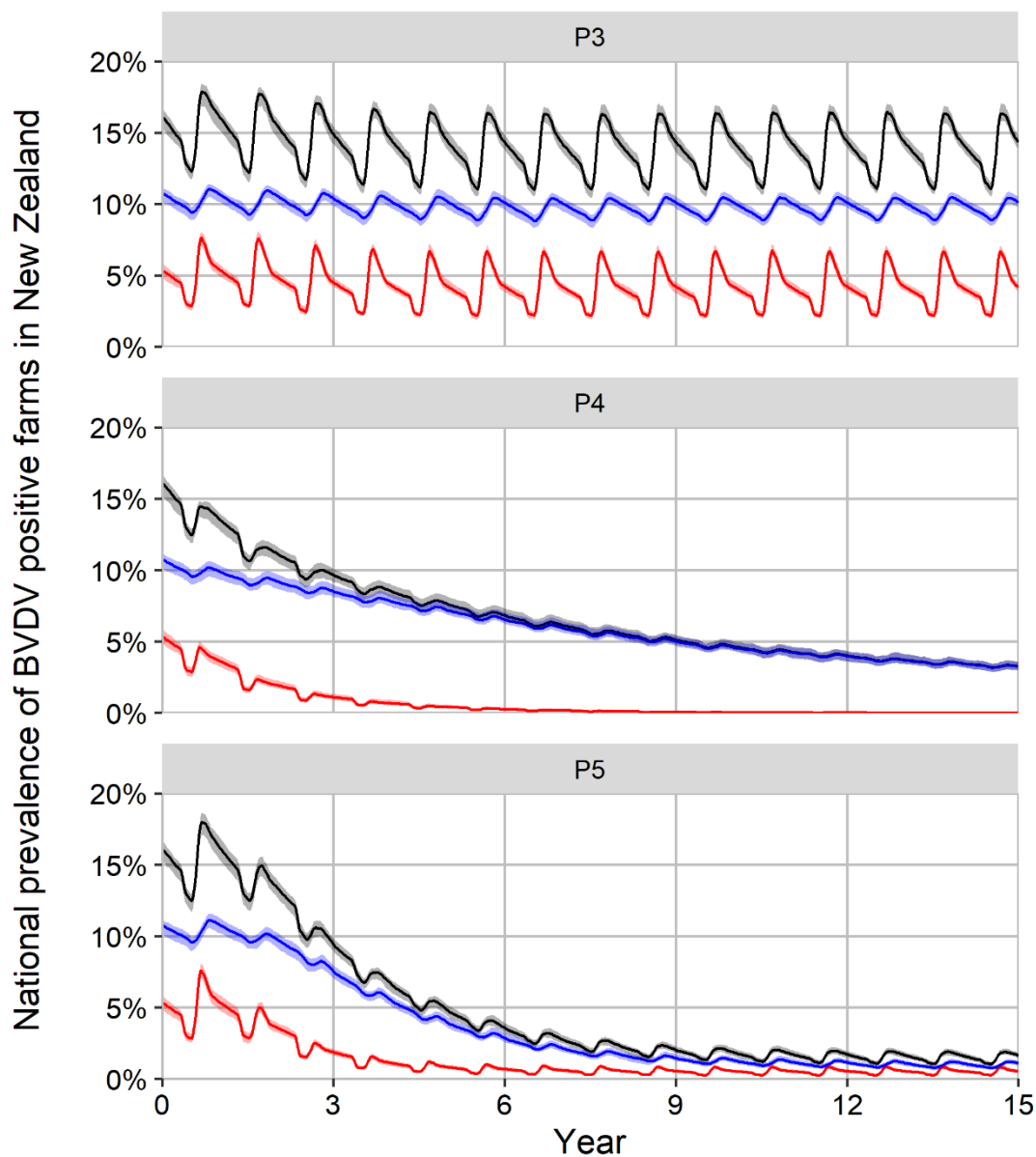


Figure 6. 5. National prevalence of BVDV positive farms (black) with respect to implementing annual screening test followed by a PI hunt (P3), double fencing (P4), and vaccination of breeding animals with annual boosters (P5) for both dairy and beef breeding farms. Red and blue lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively.

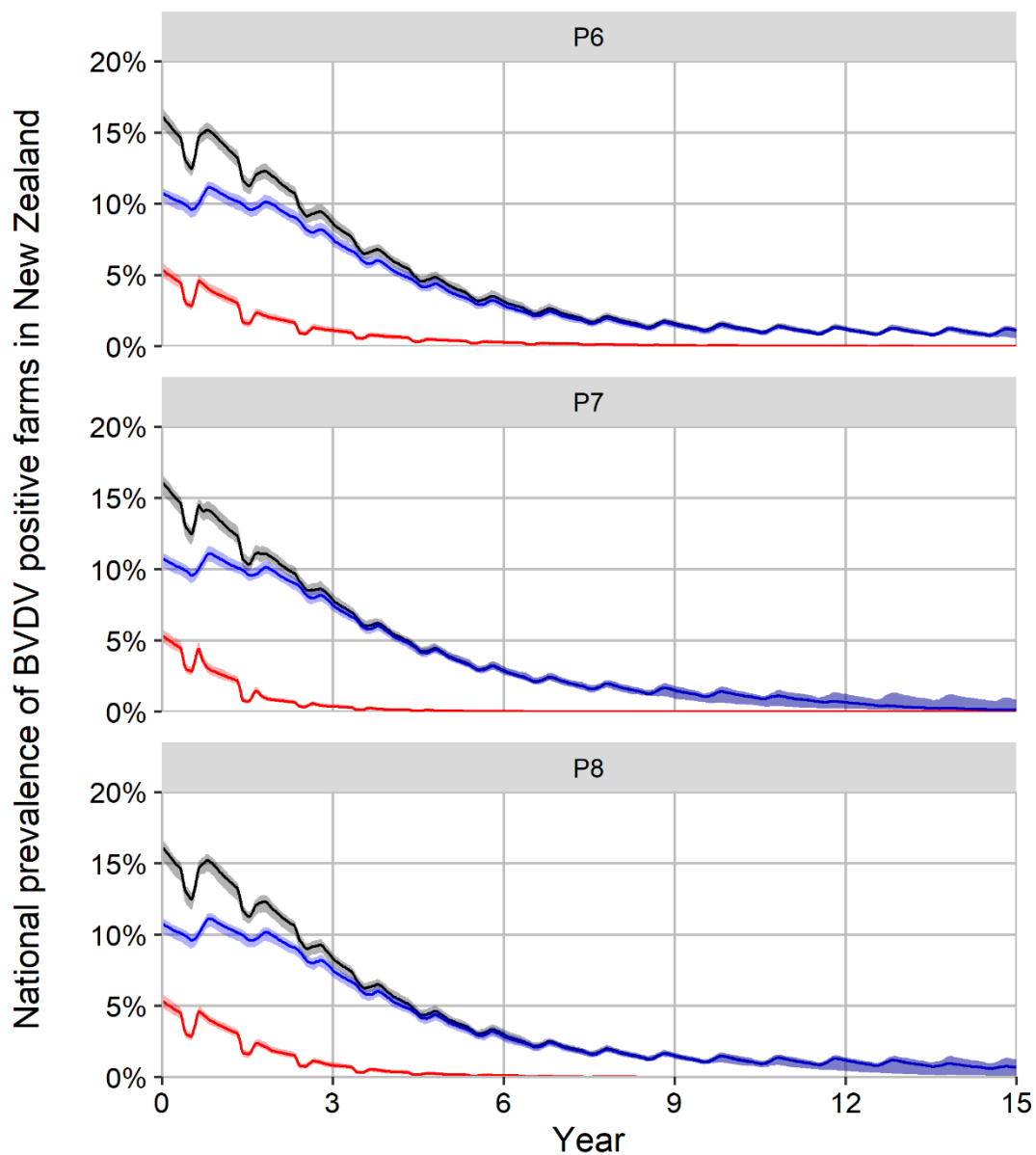


Figure 6. 6. National prevalence of BVDV positive farms (black) with respect to implementing double fencing (P6), annual testing and culling of breeding female calves with double fencing (P7), and annual screening test followed by a PI hunt with double fencing (P8) for dairy farms. For all strategies, vaccination of breeding animals with annual boosters was implemented for beef farms. Red and blue lines indicate the contribution of the dairy and beef industries, respectively, to the national BVDV prevalence. Solid line and shaded area indicate the median and 95% prediction interval, respectively.

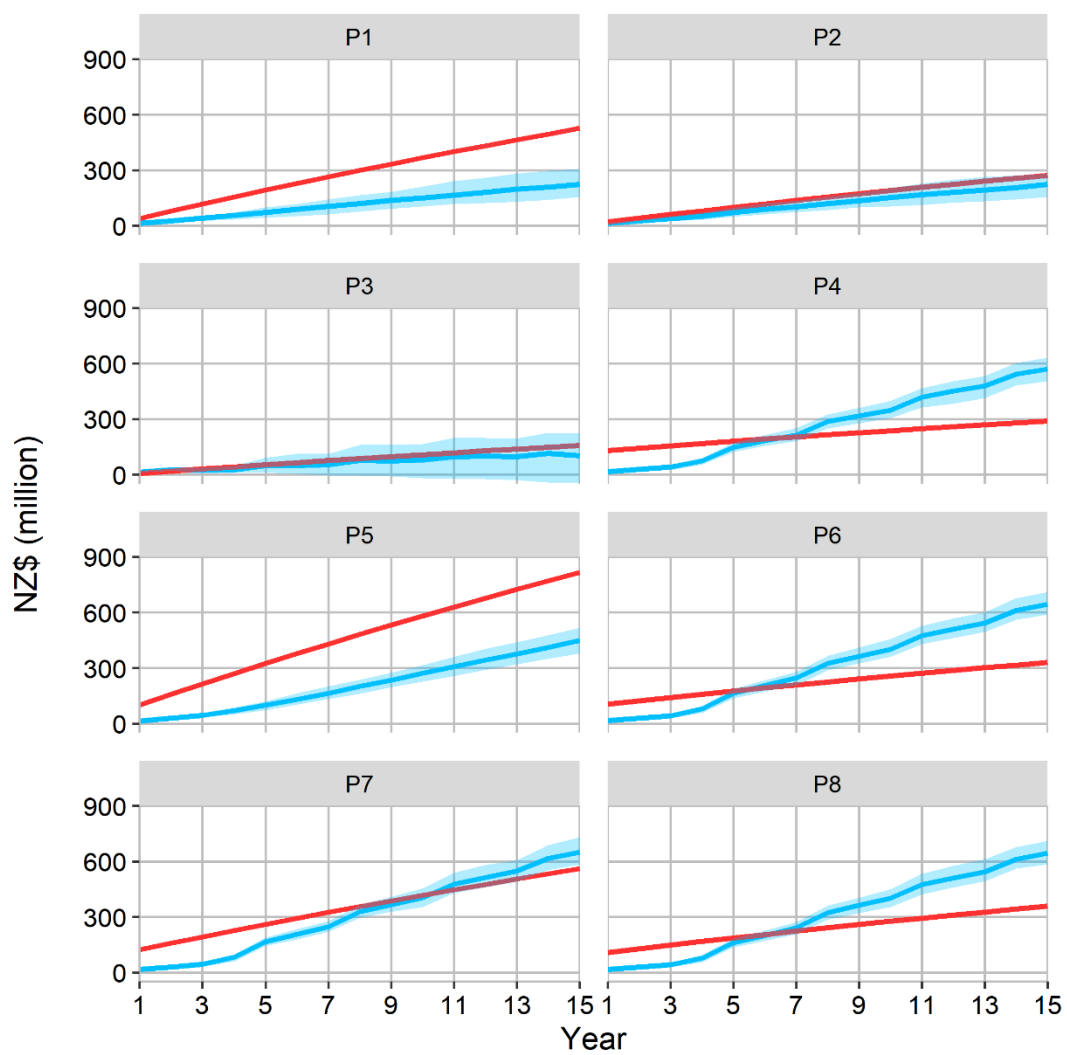


Figure 6. 7. Accumulated benefit (blue) and BVDV control cost (red) for different BVDV elimination programmes over time. Solid line and shaded area indicate median and 95% prediction interval, respectively.

6.4.2. Cost-benefit analysis of national BVDV elimination programmes

The accumulated benefit and BVDV control cost over time is shown in Figure 6.7 and Table 6.4. Implementing P7 resulted in the highest benefit with the median (95% prediction interval (PrI)) accumulated benefit estimated to be NZ\$ 650.3 million (NZ\$ 583.6 to 731.1 million) over a 15-year period. P5 was the most expensive BVDV elimination programme with the median cost of NZ\$ 815.5 million (95% PrI: NZ\$ 815.3 to 815.7 million).

The benefit/cost ratio of different elimination programmes is provided in Table 6.4. At the end of the simulation period, BVDV elimination programmes of P4, P6, P7, and P8 incurred a significantly larger benefit than cost with the median (95% PrI) benefit/cost ratio of 1.97 (1.75 ~ 2.18), 1.95 (1.78 ~ 2.14), 1.16 (1.04 ~ 1.30), and 1.80 (1.63 ~ 1.99), respectively. For each of these programmes, a break-even point was observed after the median (95% PrI) of 7 years (6 to 8), 6 years (5 to 8), 11 years (8 to 14), and 7 years (6 to 8), respectively, from the implementation of the elimination programme.

Table 6. 4. Median (95% prediction interval) of accumulated benefit (Benefit), BVDV control cost (Cost), year of break-even point, and benefit/cost ratio with respect to implementing different national BVDV elimination programmes.

Programme	Benefit (million NZ\$)	Cost (million NZ\$)	Break-even (yr)	Benefit/cost ratio
P1	222.6 (155.4, 308.1)	525.9 (525.6, 526.1)	-	0.42 (0.30, 0.59)
P2	223.9 (154.3, 282.9)	272.4 (272.3, 272.5)	- **	0.82 (0.57, 1.04)
P3	103.5 (-43.6, 223.2)	158.1 (156.2, 159.6)	- **	0.65 (-0.28, 1.41)
P4	571.5 (505.9, 631.3)	289.7 *	7 (6, 8)	1.97 (1.75, 2.18)
P5	449.8 (377.9, 516.9)	815.5 (815.3, 815.7)	-	0.55 (0.46, 0.63)
P6	644.4 (588.4, 708.8)	330.6 (330.5, 330.6)	6 (5, 8)	1.95 (1.80, 2.14)
P7	650.3 (583.6, 731.2)	560.4 (560.2, 560.5)	11 (8, 14)	1.16 (1.04, 1.30)
P8	645.0 (585.0, 710.8)	357.8 (357.4, 358.3)	7 (6, 8)	1.80 (1.63, 1.99)

Key: yr, Year.

* Accumulated BVDV control cost did not vary between simulations.

** Break-even point was beyond the simulation period for the median and upper limit of 95% prediction interval.

6.5. Discussion

This study is the first to develop a stochastic individual-based simulation model of BVDV transmission dynamics within and between New Zealand cattle farms to evaluate the cost-effectiveness of implementing a national BVDV elimination programme. The results not only highlight the important features of the epidemiology of BVDV in New Zealand, but also identify several possible approaches that can be used to eliminate BVDV with significant economic net benefits to the dairy and beef industries.

The model results highlighted key differences between the New Zealand dairy and beef industries with respect to how BVDV transmission is maintained within a farm. In dairy farms, BVDV was effectively controlled by preventing local spread via double fencing, which indicates that constant re-introduction of BVDV from external sources is required for the disease to persist. This is most likely due to the unique demographic structure of New Zealand dairy herds whereby almost all male calves and any female calves not being kept as replacements are typically sold within the first week of life, as this would eliminate any potential PI calves in these groups without the need for additional testing. It is also a common practice for replacement heifers to be grazed at off-site locations until a few months before first calving, which reduces the potential for BVDV transmission to the milking herd. It was therefore unsurprising that double fencing was an effective national control option to prevent BVDV transmission to naïve dairy stock. Although vaccination can also be used to mitigate the effects of local spread, vaccine efficacy is not perfect and it is more expensive than double-fencing since all breeding animals are required for annual boosting (Smith et al., 2014). Furthermore, double-fencing will effectively prevent local spread of other infectious diseases, such as Leptospirosis, *Mycoplasma bovis*, or Johne's disease, as well, and it will likely lead to greater long-term benefits than predicted in this study.

For the New Zealand beef industry, neither double fencing nor movement restrictions were sufficient to eliminate BVDV. We conducted a *post hoc* analysis about the cause of

ineffectiveness of double fencing on reducing the number of BVDV positive beef breeding farms (see Section 6 in Appendix 4), which suggested that larger herds (> 500 assigned MA herd size) were more likely to become endemic with BVDV. This is not surprising given that beef calves typically remain with their dams until weaning (6 ~ 7 months of age) and therefore any PI calves born into the herd have the potential to expose naïve dams during the subsequent breeding season, particular in large extensive herds where BVDV may not spread rapidly enough to confer adequate herd immunity. Lower rates of BVDV self-clearance have been reported in other studies as well (Lindberg & Houe, 2005). Given that it is logistically challenging for beef farmers to collect samples from calves at-foot to remove PI animals in a timely fashion (Han, Weir, et al., 2018), vaccination is likely to be the most practical and effective approach to control BVDV in the New Zealand beef industry. However, it is still important to prevent re-introduction of BVDV through contacts with neighbouring farms, especially since BVDV vaccines available in New Zealand have shown limited efficacy to prevent fetal infections (Packianathan et al., 2017).

Across the different national elimination programmes examined in this study, implementing double fencing for dairy farms and vaccination for beef breeding farms (P6) substantially reduced the national prevalence of BVDV positive farms. However, it should be noted that the prevalence plateaued after approximately 10 years of control and additional annual tests of breeding calves (P7) or annual herd-level screening tests (P8) on dairy farms was required for further reduction of national BVDV prevalence. This finding was related to the dynamics in the assumed background prevalence of PI animals, which contributed to local transmission in the simulation model. The background prevalence declined more rapidly when dairy breeding herds were tested or screened for BVDV because fewer non-breeding dairy PI calves were moved to beef fattening or hobby farms which would then contribute to the risk of local spread to beef breeding farms. This “*spill-over*” like phenomenon of BVDV transmission from the dairy industry to the beef sector indicates the necessity of additional BVDV testing or screening on dairy farms to eliminate

BVDV in New Zealand. Although it was not modelled in this study, there is always a risk of transmission from the beef industry back to the dairy sector through local spread and/or the movement of PI or TI breeding bulls. This is of significant concern since the levels of natural immunity in dairy herds have been decreasing over the past decade as more herds have taken measures to eliminate active BVDV infections (Gates, Han, et al., 2019; Thobokwe et al., 2004). Therefore, the model results highlight the need for both industries to work together to effectively control BVDV.

Based on the current situation in New Zealand, we therefore believe that either P7 or P8 would offer the most cost-effective and practical framework for successful elimination of BVDV in New Zealand. However, it should be noted that neither approach was able to completely eliminate BVDV from New Zealand within a time period of 15 years. This was due to the imperfect BVDV vaccine efficacy resulting in the continued creation of PI animals in beef herds. This highlights the need to view an elimination programme as a dynamic process where regulations may need to change as the epidemiology changes. For instance, the BVDV control programme in Switzerland required all new-born calves to be tested against BVDV Ag at the beginning, then this changed to Ab based screening tests to monitor BVDV incidences as the national prevalence gradually decreased (Stalder et al., 2016). Given the BVDV transmission model in a national scale of New Zealand has been developed in this study, different elimination programmes that constantly changing over time with respect to the epidemiological situation of BVDV in individual farms can be examined in future researches. It would even be possible to switch the programme to a more risk-based elimination programme similar to the approach used for bovine tuberculosis control in New Zealand (Hidano et al., 2016) whereby farms with a higher risk of BVDV introductions are targeted for more intensive surveillance and control.

We acknowledge several limitations that might affect the validity of this study. First, BVDV simulation of this study was based on poor-quality data about the number, boundary, and location of cattle farms in New Zealand. This is most likely to affect the estimates of

between-farm transmission through fence-line contacts since it is difficult to predict which farms are actually in contact with which farms. Interestingly, a recent simulation study suggested that the impact of foot-and-mouth disease in the New Zealand cattle industries could be inaccurately estimated if the variations of herd sizes and farms locations are ignored (Van Andel et al., 2018). Although there has been no study to support the impact of this information on the simulation of BVDV in New Zealand, BVDV modelling based on the inferred population in this study would have altered the contribution of different transmission routes and eventually the efficacy of different elimination programmes. Complete data for cattle farms as well as individual cattle on the farms would be more beneficial to simulate BVDV in New Zealand more accurately, however, currently such data do not exist.

Secondly, we could not accurately represent the management practices and demographic structure of individual herds on the model. For example, the PSM and thus PSC on a New Zealand cattle farm is highly affected by grass growth patterns which is strongly associated with the farm's geographical location (Anonymous, 2018a). However, in this study, we could not incorporate this association to the simulated farms due to the lack of data or compatibility between NAIT data and other data sources with management details (Jewell et al., 2016). Also, it was impossible to accurately incorporate the impact of off-site grazing since there was no data about which dairy farms were sending heifers to which grazing locations where which herds the replacement heifers may have been in contact with. These limitations would likely lead to the oversimplification of spatial heterogeneity in between-farm BVDV transmission. Therefore, it is crucial to establish a robust NAIT system that is compatible with existing data sources to facilitate accurate tracking of the disease status of individual animals and herds.

Also, the national model was calibrated based on the observed prevalence of BVDV positive farms in the dairy sector only, since there is no accurate information about the prevalence of BVDV positive beef farms due to the poor uptake of BVDV testing (Gates,

Evans, et al., 2019). Although it may indicate that parameters related to BVDV transmission in this study were incorrectly calibrated, it is worth noting that the value of calibrated parameters, such as between-farm transmission parameter or distance threshold to define adjacent farms, were similar to the values reported in other BVDV simulation studies (Ezanno et al., 2007; Qi et al., 2019).

The estimated efficacy of BVDV elimination programmes in this study was also based on an assumption that all farmers on New Zealand dairy and beef breeding farms implemented and continually complied to the programmes. In reality, however, the compliance of farmers to a disease control scheme depends on a variety of social and economic factors (Devitt et al., 2014; Kristensen & Jakobsen, 2011; McAloon et al., 2017). Simply mandating New Zealand cattle farmers to implement BVDV elimination programmes legally would not facilitate their engagement considering the current poor compliance to the NAIT system (Edge & Kavalali, 2018). Also, a recent study reported that farmers would show a risk compensation behaviour, that is the adoption of behaviours detrimental to farm biosecurity in response to the implementation of control measures (Gareth et al., 2019). This suggests that, for example, a dairy farmer who implemented double fencing might be less likely to conduct BVDV testing or screening. Moreover, scientific demonstration of the cost-effectiveness of BVDV control in New Zealand may not be sufficient to elicit supportive actions from stakeholders or industry bodies. Therefore, future studies should focus more on social science to better understand farmers' and stakeholders' acceptance of different biosecurity measures. A realistic and accurate estimation of the efficacy of BVDV elimination programmes in New Zealand will be achieved once industry compliance with the programmes is incorporated into the national BVDV simulation model.

6.6. Conclusions

Using a stochastic individual-based metapopulation model that simulated BVDV transmission dynamics in the New Zealand cattle industries, we revealed that local spread from PI animals in neighbouring farms is the most important transmission route for the industries to maintain the endemic state. Also, we demonstrated that substantial reduction of national BVDV prevalence is achievable with implementing elimination programmes based on double fencing for dairy farms and annual vaccination for beef breeding farms, however, more rigorous control options would be needed to completely eliminate the disease as national prevalence declines. Although this study suggested that elimination of BVDV is both technically feasible and cost-effective for the New Zealand cattle industries, discussions with farmers and stakeholders will be required to facilitate their compliance to national BVDV elimination programmes in order to successfully eliminate the disease.

6.7. Acknowledgements

We are grateful to New Zealand beef farmers, veterinarians, and laboratory technicians who participated in the “BVD Free New Zealand” project.

Chapter 7

General discussion

7.1. Overview of key findings

This thesis evaluated the cost-effectiveness of national bovine viral diarrhoea virus (BVDV) elimination programmes in New Zealand by addressing the knowledge gaps around the epidemiology and economics of BVDV and its control in the New Zealand cattle industries.

One of the previously recognised knowledge gaps was around the risk factors for BVDV infection in pasture-based New Zealand cattle herds. In Chapter 3, a risk factor analysis using data from 304 herds participating in a national cross-sectional study revealed that local spread between neighbouring farms as well as onto-farm movements of breeding heifers/cows are the key transmission routes of BVDV in the New Zealand cattle industries.

Another epidemiological knowledge gap was the speed at which BVDV spread within extensively grazed beef herds in New Zealand. In Chapter 4, the within-herd BVDV transmission rate was estimated using longitudinal seroconversion data from nine New Zealand beef breeding farms. The transmission rate was inferred to be 0.11 per persistently infected (PI) animal per day, suggesting that the spread of BVDV in New Zealand beef herds is slower than in dairy herds.

Based on the findings from the previous chapters, BVDV transmission models for typical New Zealand dairy and beef breeding farms were developed in Chapter 5. Using this modelling approach, the direct losses due to a BVDV outbreak on a naïve dairy and beef farm were estimated at NZ\$ 22.22 and NZ\$ 41.19 per mixed-age cow per year, respectively. Also, it was revealed that annual testing of breeding calves to cull identified PI animals for dairy farms and annual vaccination of breeding animals for beef farms would

be the most economically beneficial strategies for controlling a BVDV outbreak on New Zealand dairy farms and beef farms, respectively.

In Chapter 6, the BVDV simulation modelling approach was scaled up to the national level using data from the national animal identification and tracing (NAIT) system to infer the demographics of the New Zealand cattle industries as well as between-farm BVDV transmission through animal movements and neighbourhood contacts. This study highlighted that the endemic state of BVDV infection in New Zealand is primarily maintained through local spread between cattle herds on adjacent farms. This study also demonstrated that a substantial reduction in national BVDV prevalence could be achieved in a cost-effective manner by requiring all dairy farms for double fencing of boundaries with either annual testing and culling or annual herd-level screening and all beef breeding farms for annual vaccination of every female breeding animal.

7.2. Risk factors for BVDV infection in New Zealand cattle farms

The risk factor analysis for BVDV infection in this thesis (chapter 3) revealed that having close contacts with animals on neighbouring farms significantly increased the risk of BVDV infection on New Zealand cattle farms. Local spread through nose-to-nose contact over fence-lines has been identified as a common transmission route of BVDV in many countries (Bitsch et al., 2000; Valle et al., 1999) and a previous study from New Zealand also showed that over-the-fence contacts with cattle on neighbouring farms resulted in increased BVDV antibody (Ab) titres in bulk tank milk (BTM) samples from dairy herds (Weir et al., 2016). These findings were corroborated through the national simulation model (chapter 6), which indicated that local spread is the primary reason for BVDV remains endemic in the New Zealand dairy industry. A similar finding has been reported in France where an estimated 93.1% of BVDV incidences at a farm-level were attributable to neighbourhood contacts (Qi et al., 2019). The importance of biosecurity

measures to prevent local spread has been repeatedly highlighted by Scandinavian countries where BVDV has been eliminated (Ståhl & Alenius, 2012). Given that double fencing of shared boundaries can prevent the spread of many other infectious diseases, this control measure should be strongly considered for inclusion as part of any national disease control programme in New Zealand.

The study also identified onto-farm movements of breeding heifers/cows as a significant risk factor for BVDV infection in New Zealand cattle herds, which is similar to findings from previous studies. Using RT-PCR on BTM from New Zealand dairy herds, Weir et al., (2016) showed that herds had purchased cows with unknown BVDV status were 2.2 times more likely to have active circulation of BVDV. In beef farms, W. Cuttance & Cuttance, (2014) reported that purchasing replacement heifers significantly increased the odds of BVDV infection although the study suffered from a lack of precision. The movement of animals between cattle farms is a well-established risk factor for BVDV infection (Valle et al., 1999). In particular, there is a significant risk that purchasing pregnant female breeding animals with unknown BVDV infection status would result in the introduction of BVDV through the birth of a PI calf. Although it is possible to assess the BVDV infection status of a fetus prior to the purchase of pregnant animals, it is very labour-intensive and may induce abortion as it requires fetal fluid for testing (Callan et al., 2002). Another approach to diagnose Trojan dams (i.e. non-PI dams that carrying PI fetus) is to use BVDV Ab ELISA test, however, the test is valid only after 7 months of gestation period with moderate specificity (0.7) (Lindberg et al., 2001). Therefore, national control programmes that restrict the movement of cattle with unknown BVDV status are likely to be practical and effective in reducing the between-farm BVDV transmission in New Zealand. In cases where movement restrictions are not feasible, the calves of any purchased pregnant dams should be tested for BVDV as early as possible after birth.

In the New Zealand dairy industry, it is common practise for farmers to send replacement heifers to off-site locations for grazing and breeding before they return to the

mixed-age cow herd just prior to delivering their first calf (i.e. off-site grazing). This carries substantial risk of introducing BVDV to the milking herd through the birth of PI calves if naïve heifers were exposed to BVDV during the risk period of pregnancy, particularly if they were unvaccinated or if there was poor biosecurity at the off-site location. It has previously been reported that BVDV BTM Ab titres of New Zealand dairy farms that grazed heifers off-site were higher than the ones that raised heifers on farm (Weir et al., 2016), suggesting that there is greater exposure to BVDV at off-site locations. A study from Switzerland also reported a similar finding that grazing animals at a distant location increased the risk of BVDV infection (Presi et al., 2011). Although it was not possible to confirm that the off-farm movements of either replacement heifers or cows increased the risk of BVDV infection in Chapter 3, off-site grazing still poses a risk of BVDV introduction for New Zealand dairy herds given the endemic status of BVDV in New Zealand. Considering the current high background prevalence of BVDV in New Zealand, vaccinating heifers before off-site grazing and testing the calves of the heifers would be recommended.

Overall, this study suggests that local spread over fence-lines and the movements of female breeding animals with unknown BVDV infection status should be targeted to eliminate BVDV in New Zealand cattle farms. However, the Bayesian network (BN) approach in Chapter 3 revealed that onto- and off- farm movements of heifers are frequent in New Zealand dairy farms whereas beef farms were less prone to these risk factors. The BN approach also discovered that New Zealand beef farmers had lower awareness of BVDV, which may make it more difficult to convince farmers to conduct BVDV control measures, such as annual BVDV testing to identify and cull PI animals, given the logistical challenges of routinely gathering animals for identifying and sampling (Gates, Evans, et al., 2019). These findings indicate that movement restrictions that require a rigorous process of testing individual animals may be more suitable and effective in the dairy industry. This speculation was supported by the national BVDV simulation study (chapter

6), which showed that implementing national movement restriction reduced up to 38.0% of BVDV positive farms in the dairy sector yet only resulted in a 4.6% decrease in BVDV positive farms in the beef industry. Conclusively, these findings imply that BVDV transmission between cattle farms in New Zealand would be better controlled with the application of industry-specific control measures.

7.3. BVDV dynamics in extensively grazed herds

One of the significant contributions of this thesis was estimating the rate of BVDV transmission in extensive beef breeding herds by using longitudinal data from a field study to measure the seroconversion of first-calf heifers over the mating and early gestation period (chapter 4). Many research groups have developed BVDV simulation models to suit their own epidemiological circumstances, and Viet et al., (2004) reviewed the BVDV dynamics of different modelling studies by comparing the number of animals that seroconverted during a year following the introduction of a single PI animal in a herd of 100 susceptible animals. The authors found that 30 to 100 animals seroconverted in that year with most herds in the study being intensive dairy herds. The estimate of 33 animals infected per PI animal per year for New Zealand beef herds (chapter 4) confirms the general belief of that BVDV transmission in extensively grazed beef herds is slower than in herds that are more intensively managed and grazed.

An important implication of this low transmission rate is that the presence of PI animals in extensively grazed beef herds is not necessarily sufficient to ensure that naïve animals will be naturally exposed to BVDV prior to the start of mating and therefore protected against producing new PI calves. A similar finding was reported in an extensively grazed dairy herd where approximately 20% of replacement heifers remained susceptible to BVDV even after being directly co-grazed with PI animals for up to 600 days (Thompson, 2005). This strongly suggests that using PI animals as a “*natural*” BVDV vaccination

source (i.e. purposely exposing PI animals to susceptible animals to replicate the protective effect of BVDV vaccination) is likely to fail in extensively grazed cattle herds in New Zealand.

Another implication of this finding with respect to BVDV control in New Zealand is that detecting a recent BVDV introduction in a naïve beef herd by conducting a spot test (i.e. BVDV Ab ELISA using a pooled serum sample of 10 to 15 animals) may have poor sensitivity if there has not been sufficient time for enough animals to be exposed to the virus. In European countries where BVDV has been systematically controlled, spot tests using blood samples from 10 to 15 youngstock are often used as an inexpensive screening test to make inferences about farm-level BVDV infection status to determine if a PI hunt is warranted (Houe et al., 2006). Given the slower BVDV dynamics in extensively grazed beef farms in New Zealand, the application of a similar spot test is unlikely to be a reliable indicator of BVDV infection status at a herd-level, so future studies should be conducted to explore the sensitivity and specificity of using Ab-based spot test to detect active circulation of BVDV.

7.4. Economics of BVDV and its control

In New Zealand, the direct losses from a BVDV outbreak on typical dairy and beef breeding farms were estimated to be NZ\$ 22.22 and NZ\$ 41.19 per mixed-age cow per year, respectively (chapter 5). These estimated losses were relatively lower than estimates from other countries, which have ranged from NZ\$ 4 to 4,512 per dairy animal and from NZ\$ 41 to 504 per beef animal (Richter et al., 2017). Although the economic impacts of BVDV depend on a variety of factors, such as the value of cattle or cattle products, production level, and management features, these findings imply that the current impacts of BVDV infection in the New Zealand cattle industries may be less than in other countries due to (1) the unique management characteristics of the dairy industry; a narrow seasonal

breeding period, selling/culling most calves at approximately 4 days old, grazing heifers off-site, and lower average milk yields per cow, and (2) a slower BVDV transmission rate caused by extensive grazing with a lack of frequent gathering events in the beef sector (chapter 4). However, it is worth noting that most of the previous studies estimated the economic impacts by comparing the difference in average production levels or reproductive performances between herds with and without BVDV. It is likely that herds with BVDV have other infectious disease or management issues on farm that could be contributing to their poor performance. Given that the estimation in Chapter 5 was based on the direct impact of BVDV at an animal-level, it is likely a reasonable and conservative estimate of the economic impacts of BVDV in New Zealand.

The within-herd simulation models in Chapter 5 revealed that control of BVDV on a naïve dairy farm was economically beneficial for some, but not all combinations of available control measures. The highest benefit/cost ratio of 2.02 occurred with implementing annual testing of female breeding calves to identify and remove PI animals. Adding double fencing to the annual testing also incurred marginally greater benefit than the cost of control (benefit/cost ratio of 1.02). Other strategies were not economically beneficial, which was in contrast to a study by Reichel et al., (2008) who suggested that implementing any type of control measures incurred more benefits than costs for a typical New Zealand dairy farm. However, the previous study likely overestimated the economic impacts of BVDV infection as it assumed the continuous presence of PI animals in the herd for the 10 years of the study period. Considering the high self-clearance rate of BVDV infection in New Zealand dairy herds (Gates, Han, et al., 2019), it is more likely that not all BVDV control options would be economical for New Zealand dairy farms.

Several strategies that based on the annual vaccination of breeding female animals were identified as being cost-effective to control a BVDV outbreak on a typical New Zealand beef breeding farm (chapter 5). Interestingly, implementing additional control measures on top of vaccination, such as annual testing of breeding calves or performing annual screening

tests to determine if a PI hunt was warranted generated only limited benefits. This was likely due to the model assumption that testing or screening was conducted around when calves were weaned. Since calves were co-grazed with dams for 6 to 7 months before weaning in the model to reflect common practises in the beef industry, there was ample opportunity for susceptible dams to be exposed to new-born PI calves during the risk period of pregnancy for generating a new PI fetus. However, the time of weaning is the most logistically convenient time for beef farmers to conduct sampling especially given the risk of injuries of calves at-foot would be an obstacle to conduct early testing and culling of PI beef calves (Gates, Evans, et al., 2019; Han, Weir, et al., 2018). These results indicate that vaccination would be the most practical and effective approach to control BVDV in the New Zealand beef industry unless farmers are willing to test and cull PI calves prior to the start of the next mating period.

7.5. National BVDV elimination programmes

In the national scale model (chapter 6), any elimination programmes that required all dairy farms to double-fence shared boundaries with either annual testing and culling of all breeding calves or annual herd-level screening tests followed by a PI hunt and all beef herds to vaccinate every female breeding animal before the mating period effectively reduced national BVDV prevalence in New Zealand with the median of benefit/cost ratios ranging from 1.16 to 1.80 and the time until break-even ranging from 7 to 11 years (median). Interestingly, this is similar to the estimated break-even points of 2 to 15 years in other European countries with either regional or national elimination programmes (Richter et al., 2017). This highlights that eliminating BVDV on a national scale in New Zealand is technically feasible with a similar level of cost-efficiency compared with other countries even though the economic impact of BVDV infection on individual farms is generally lower in New Zealand.

A challenge for national BVDV elimination in New Zealand is the strong linkage between the dairy and beef industries due to the estimated 0.8 million dairy calves that are sold annually to the beef industry for fattening purpose (*per. comm.* Steve Morris). Considering the cost and labour of BVDV antigen (Ag) testing to identify PI calves, it is uncommon for New Zealand dairy farmers to test calves before selling them to beef farms, which means that a substantial number of PI calves are likely moved from the dairy to beef fattening farms. Given that the persistence of BVDV in the dairy industry was mainly due to local spread from PI animals on neighbouring farms (chapter 6), this “*spill-over*” like phenomenon of PI animals poses a great risk of BVDV transmission from the beef sector back to the dairy industry via neighbourhood contacts. This highlights the need for the dairy and beef industries to work together simultaneously to eliminate BVDV in New Zealand.

Due to the imperfect efficacy of vaccination programmes implemented in beef breeding herds, the simulation model in Chapter 6 predicted that none of the national elimination programmes that were examined in the model would completely eliminate BVDV from New Zealand within 15 years although the national prevalence could still be substantially reduced. This indicates that any BVDV elimination programmes implemented in New Zealand will likely need to evolve over time as the epidemiologic situation changes. In Ireland, for example, a recent study suggested a legislative change to mandate farmers to cull PI animals sooner than initially required to achieve further reduction in the national BVDV prevalence with an attractive benefit/cost (Thulke et al., 2018). One important point to note is that the cost-effectiveness of BVDV elimination programmes often decreases substantially during the final phase of BVDV elimination due to the high costs of rigorous searching for any remaining PI animals in the small number of actively infected herds. As an alternative, the implementation of risk-based approaches that selectively target farms with a high anticipated risk of BVDV introduction may prove more beneficial. However, this approach has never been applied in other countries possibly due to the uncertainties

around accurately predicting which farms may be infected, so questions still remain about the efficacy of such programmes.

7.6. Methodological limitations

As with many research studies, there were limitations in both the data and methodological approaches used in this thesis although every effort was made to ensure accuracy and validity of the results.

The study in Chapter 3 was based on secondary data collected from a convenience sample of 304 New Zealand cattle farms, therefore, the study was inherently influenced by selection bias. However, although the magnitude and direction of the bias is unknown, the selection bias likely had a limited impact given that the estimated prevalence of active BVDV infection on New Zealand cattle farms was consistent with those reported from other studies that had a wider sampling frame (Gates, Han, et al., 2019; Weir et al., 2016). One approach that could have been applied to adjust the selection bias is a quantitative bias analysis, which adjusts the coefficients of explanatory variables (or exposure) by considering the probability of being sampled depending on the status of exposure and outcome of cattle farms (Lash et al., 2011). However, the method is only available to the univariable approach and application to multivariable framework is still limited.

Another limitation is that farms were considered actively infected with BVDV if the S/P ratio from the pooled serum BVDV antibody (Ab) ELISA was > 0.75 . This cut-off value was based on a previous study showing that it was extremely unlikely for herds with an S/P ratio less than 0.75 to have any PI animals present on farm (Hill, McCoy, et al., 2010). However, the results from the panel study in beef herds provided empirical evidence of BVDV seroconversion occurring in herds with low S/P ratios (Gates, Evans, et al., 2019), which indicates that classifying a farm's active BVDV infection status using BVDV Ab ELISA on a pooled sample is not completely accurate. Unfortunately, given the nature of

secondary data, it was not possible to confirm whether BVDV was actively circulating on the 304 sampled farms. Considering that approximately half of dairy farms were categorised as actively infected with BVDV in the study while it has been recently estimated as approximately 7% (Gates, Han, et al., 2019), it would indicate that the large number of farms with active BVDV infection in Chapter 3 was false positive, which possibly drove most of coefficients of explanatory variables into the null. The survey design could affect the association between the explanatory and outcome variables even worse, since the survey captured the risk of BVDV introduction or incidence during the past 5 years while the outcome measured the possible evidence of current BVDV persistence within farms. Future studies on risk factor analysis for BVDV infection should be based on more evident outcomes such as the confirmation of BVDV Ag or presence of PI animals.

In Chapter 4, an approximate Bayesian computation-sequential Monte Carlo method was used to infer the value of model parameters that related to BVDV transmission in extensively grazed beef herds in New Zealand (e.g. BVDV transmission rate, initial proportion of BVDV seropositive animals, proportion of introduced PI animals, and day of PI animals being introduced) with empirical data on the seroconversion of first-calf heifers between two sampling events. With this relatively limited data, particularly without knowing how many PI animals were present in the herds, it was not possible to accurately estimate all unknown model parameters. Future studies should track the serological status of the whole herd across more sampling points to better estimate seroconversion rates. However, as previously highlighted, it is logistically challenging and expensive to perform intensive sampling in extensive New Zealand beef farms since cattle are infrequently handled.

The stochastic individual-based BVDV simulation models developed in Chapter 5 were based on many assumptions about the demographic structure of an average dairy or beef farm in New Zealand as well as assumptions around the impact and cost of BVDV outbreaks. For example, the model assumed that beef farmers would only conduct annual

screening tests of breeding calves at weaning, which was not a cost-effective strategy since the removal of PI calves was too late to break the BVDV transmission cycle. It is possible that this strategy would have been more cost-effective if farmers were willing to yard and test calves prior to the start of the next mating period. The model results of economic impacts were also highly dependent on both the number of Trojan cows (i.e. non-PI dams carrying a PI fetus) that were introduced to seed the outbreak and the mortality rate of PI animals, and the value of these parameters used in this study may not accurately reflect the true epidemiological situation for all individual herds in New Zealand. As another future direction for research, it would be interesting to use longitudinal data about production performance, such as milk yield and live-weight gain, of individual animals with their BVDV infection status over time and then fit the models to more accurately predict the economic benefits of controlling BVDV.

The stochastic individual-based metapopulation model developed for Chapter 6 was significantly limited by the lack of complete and accurate data on the demographics of individual cattle and cattle farms in the NAIT system as well as the lack of information on the true BVDV status of individual cattle. Given that the simulation model was based on inferred demographic structures, it is likely that the dynamics of BVDV transmission in the New Zealand cattle population differ in the real world. Furthermore, the cattle trade movements for the model were also simulated from NAIT data, which has known issues with underreporting of movements, particularly those involving the temporary movement of animals to off-site locations (Edge & Kavalali, 2018). The contribution of animal movements to between-farm BVDV transmission may therefore be underestimated in the national model, leading to a further underestimation in the cost-effectiveness of movement restriction as a strategy to control BVDV in New Zealand. However, it is worth noting that less than 9% of PI animals generated in Irish herds were attributable to cattle movements (Reardon et al., 2018), so it is possible that the impact of underestimated animal movements on national BVDV transmission in New Zealand might also be negligible. Regardless,

future effort should be made to establish a database that effectively collects the demographics and movement of individual cattle and records their BVDV infection status in order to better understand between-farm BVDV transmission in the country.

7.7. Future work: feasibility of BVDV elimination in New Zealand

Largely due to efforts by the BVD Steering Committee in providing extension resources to veterinarians, there has been significant progress towards increasing awareness of BVDV and the uptake of control measures in the New Zealand cattle industries (Stewart, 2013). It has led to the reduction of overall prevalence of farms with active BVDV infection over the last 10 years (Gates, Han, et al., 2019; Thobokwe et al., 2004). Once BVDV is eliminated from New Zealand, it would be relatively easy for the country to maintain the freedom given the remote geographic location with strict biosecurity measures at the border.

However, the readiness of BVDV elimination for the New Zealand cattle industries is still in question. According to Dowdle, (1999), three preconditions — (1) an effective measure to intervene the transmission of the host, (2) practical diagnostic tests with high sensitivity and specificity, and (3) a lack of other reservoir hosts that amplify the disease in the environment — are crucial to eliminate a disease. Although this thesis has generated additional insights into the cost-effectiveness of different BVDV elimination programmes designed with commercially available and reliable diagnostic tests to intervene the transmission of BVDV in New Zealand, there are still knowledge gaps that should be targeted as a part of future research for the New Zealand cattle industry to meet these preconditions.

7.7.1. Accurate data

First, there is a general lack of accurate data about the demographics of New Zealand cattle, particularly in the beef sector. In order to support future disease control efforts, data recording systems need to capture the current list or number of animals within each farm,

number of active or commercial beef farms, boundary or exact location of cattle farms, and the production type of cattle farms (e.g. beef breeding, beef fattening, or bull breeding). Moreover, animal movements between different premises are not properly captured due to the lack of compliance of farmers, and it is obscure how many animals are grazed off-site at which locations, or how often animals owned by share-milkers are moved between different premises. Also, only the overall production data of the country (i.e. national milk solid production per year, total weight of exported beef meat, or total number of slaughtered animals per month) is available, and information about the production performance of individual herds is either not recorded or recorded without tracking their infection status of BVDV or other diseases. Although there are some platforms (e.g. AgriBase, FarmsOnLine, MINDA, or NAIT) that have the capacity to capture some of this information, the data collected via these systems are not generally up-to-date and there is a lack of compatibility between the different databases that impedes data aggregation (Jewell et al., 2016). Poor demographic data quality is known to have a significant impact on simulation model results. For instance, Van Andel et al., (2018) showed that disease modelling based on inaccurate demographic information led to different estimates of the length and severity of foot-and-mouth disease epidemics in New Zealand compared with models based on a “*true*” dataset. This highlights the necessity of establishing a database system with accurate demographic and production information and disease status to track epidemiological situation over time.

7.7.2. Social acceptability

Although the simulation models predicted that BVDV control would be cost-effective at both the herd and national level, gaining social acceptance for implementing a national BVDV control is likely to be a significant hurdle. Previous studies have shown that a variety of factors, including farmers’ knowledge about the disease, herd size, market-driven economic premiums, or social and cultural significance (e.g. recognised as “good farmer” by others), are complexly associated with their decision to participate in control programmes (Garforth et al., 2013; Kristensen & Jakobsen, 2011; McAloon et al., 2017;

Nöremark et al., 2010). In New Zealand, a recent study of beef cattle farmer opinions towards BVDV control noted that animal health concerns such as parasite infections or mineral deficiencies were higher on the list of farmer priorities than BVDV (Gates, Evans, et al., 2019). Also, this thesis showed that there was no significant correlation between New Zealand farmers' knowledge about BVDV and their implementation of biosecurity measures to prevent the disease (chapter 3). Therefore, it is unlikely that simply providing farmers and industry stakeholders with information about the economic incentives for eliminating a disease would be sufficient to make them implement any control measures (van Asseldonk et al., 2010). Further research into the social science of understanding what motivates New Zealand farmers and stakeholders to join a control programme are needed to identify more effective strategies to engage farmers, veterinarians, and industry stakeholders to implement a systematic control programme that meets their needs.

7.7.3. Governmental support

Under the current legislative framework in New Zealand, any costs related to government veterinary services for “endemic disease control and eradication” are covered by livestock owners rather than the Ministry for Primary Industries (Lawrence, 2001). Accordingly, since BVDV has been endemic since at least the 1960s (Salisbury et al., 1961), there have been no government subsidies to control BVDV in the New Zealand cattle industry. However, it has been noted in the European countries with formal BVDV control programmes that legislative and financial support for farmers from governments is essential to successful BVDV control (Bitsch et al., 2000; Lindberg et al., 2006). Farmers in Ireland receive financial compensation for culling identified PI animals on their farms (Graham et al., 2014), and subsidies from the government for testing, vaccinating, and culling animals to control BVDV may help facilitate farmer compliance with national BVDV control programmes in New Zealand.

7.7.4. Other possible limitations

Although RT-PCR of BVDV is commonly used in the commercial diagnostic laboratory, there is limited knowledge of the genotypes of endemic BVDV strains in New Zealand with the last reported study conducted in 1998 (Vilček et al., 1998). It has been proven in countries with genetic databanks of prevalent BVDV strains that molecular epidemiology is a useful tool to increase the understanding of BVDV transmission (Booth et al., 2013; Ståhl et al., 2005). By conducting phylogenetic analysis to investigate the epidemiological links between field BVDV strains collected from different cattle farms, better understanding about the important routes of BVDV transmission between New Zealand herds could be achieved.

Finally, the role of sheep in BVDV transmission in the New Zealand cattle industries also warrants further investigation, especially given the large number of New Zealand beef herds are co-grazed with sheep. Even though sheep are considered spill-over hosts of BVDV from cattle populations, a recent study found serological evidence of BVDV infection in sheep flocks that were co-grazed on New Zealand cattle farms where active BVDV circulation was highly suspected (Evans, Han, et al., 2019). Also, a case report has shown that BVDV PI lambs can be generated through close contact with PI heifers (King, 2014). Although no studies have yet demonstrated BVDV transmission from PI sheep back to naïve cattle on New Zealand farms, there has been a report in the literature of a PI ram serving as the source of BVDV circulation on some Danish cattle farms (Bitsch et al., 2000). Accounting for the epidemiological and economic impacts of BVDV in sheep could potentially make national BVDV control even more cost-effective for New Zealand cattle farms.

7.8. Conclusion

BVDV is prevalent in cattle farms in New Zealand, and its economic impacts including direct and indirect losses and current expenses for voluntary controls on the whole cattle industries are believed to be significant. To propose cost-effective national BVDV control programmes for elimination of the disease, the studies in this thesis analysed the risk factors for BVDV infection and estimated the within-herd BVDV transmission rate to better understand the epidemiology of BVDV in the New Zealand cattle industries. Also, the direct losses of a BVDV outbreak and cost-effectiveness of BVDV control in New Zealand dairy and beef breeding farms were measured to address the economic knowledge gaps around BVDV and its control in New Zealand. Finally, it was demonstrated that some control programmes could substantially reduce national BVDV prevalence with attractive benefit/cost ratios. With these findings, this thesis delivers a key message that successful BVDV control is technically feasible and economically beneficial in the New Zealand cattle industries.

Although these results support the technical feasibility of successful control of BVDV, it should be noted that, in countries where BVDV has been eliminated, the implementation of BVDV control programmes was initiated by farmers and their representative groups (Lindberg et al., 2006). Thus, elimination of BVDV in New Zealand will not be achievable without the compliance and cooperation of farmers and industry stakeholders. Eliciting support for national BVDV control programmes from farmers and industry stakeholders is likely to be a challenging task given the current low uptake of voluntary control measure among New Zealand farmers. Therefore, I sincerely hope the findings of this thesis ignite farmers, veterinarians, and stakeholders to further engage in discussions about the value and achievability of national BVDV elimination in the New Zealand farming context.

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Appendix 1

Section 1. Original survey form



See further explanation for each question on the rear of this form if required.
PLEASE CIRCLE THE APPROPRIATE ANSWER.

1. Farm ID		2. Region		7. Date sampled (dd/mm/yy)			
				/ /			
3. Type		4a. Is the primary (home) farm one continuous block of land?		8. Vet/Clinic			
DAIRY		YES NO					
BEEF only		4b. How many blocks?				5. How many adjacent farms / land with cattle?	
SHEEP & BEEF		0 1 2 3 4+				0 1 2 3 4+	
6. Over the last five years or so, in terms of cattle not bred on your farm, what classes of stock have come onto the farm							
a heifers and/or cows for the breeding/milking herd?		all/most years		occasionally/rarely never			
b bulls to use for mating?		all/most years		occasionally/rarely never			
c weaner or store cattle to on-sell or finish?		all/most years		occasionally/rarely never			
d replacement dairy heifers for grazing?		all/most years		occasionally/rarely never			
e winter/carry-over dairy cows for grazing?		all/most years		occasionally/rarely never			
Do you share cattle yards with your neighbour(s)?		all/most years		occasionally/rarely never			
9. No. in the mob?		10. In what season was the mob born?		11. In what year was the mob born?			
		spring autumn other		2013 2014			
12. Were the heifer test mob born on your farm?		all some most none unknown					
13A. Have any of the test mob been off your farm?			13B. If off the farm, approximately how long have they spent away?				
all some most none unknown			Grazing off your farm <1mth 1-3mths >3mths n/a				
14A. If mated, how were the mob mated?			14B. If bulls used, were they				
AI only Bull only AI & Bull N/A			a BVD tested?				
15. Do the mob have nose to nose contact with stock from your farm?			all some none unknown				
often occasional rarely never			b certified BVD free?				
16. Do the mob have nose to nose contact with stock from other farms?			all some none unknown				
often occasional rarely never			c BVD vaccinated?				
17. Were the dams of the test mob born on the farm?			18A. Were the dams off the home farm during mating pregnancy with the test mob?				
all most some none unknown			all most some none unknown				
19A. How were the dams of the mob mated?			18B. If off farm, (circle one or more)				
AI only Bull only AI & Bull N/A			a where did they go?				
19B. If bulls used, were they			grazing neighbours yards run-off other				
a BVD tested?			b how were they transported?				
all some none unknown			carrier farm truck trailer walked other				
b certified BVD free?							
all some none unknown							
c BVD vaccinated?							
all some none unknown							
20A. If a dairy farm, have you had a bulk BVD test done?		20B. If yes when, and what was the last result?					
YES NO							
21. How would you rate your knowledge of BVD?		22. Do you know what a PI animal is?		23. Do you know what a trojan animal is?			
0 poor to 4 excellent		YES NO		YES NO			
0 1 2 3 4+							
24. Other comments							

Section 2. Merging and re-categorising variables


We reduced the dimensionality of the imputed dataset by re-categorising responses of variables. Production type was divided into either “Dairy” or “Beef” farms. Variables of neighbouring farms and awareness of BVD were re-categorised by aggregating responses of “3” and “4+” into one group and other responses into another group. Frequencies of close animal contacts within and between farms were categorised into either “Often”, “Occasionally/rarely” or “Never”. Off-farm movements of heifers and mixed-age breeding cows were grouped in such a way to fit responses into either “None” or “Not-none”. Mating methods for both heifers and mixed-age cows were re-categorising as “Bulls used” or “Bulls not used”.

Also, for the purpose of decreasing the degree of freedom of regression models, a variable of mating method and three variables of BVD prevention measures of service bulls were merged for each two group (e.g. herds of replacement heifers and mixed-age breeding cows). Given that the variables of BVD prevention measures of service bulls were nested variables that only allowed to group of farmers responded as “Bulls used” in respective mating methods, we integrated each of mother and daughter variable set in two-fold. First, for each group, three variables of BVD tested, certified, and vaccinated service bulls were aggregated as a new variable of “BVD prevention status of bulls”. Responses of the new variables were set in a way that (1) if response of any variable for merging was “All”, then the response of the new variable became “All prevented”, (2) when all variables for merging showed “None/unknown”, then the response of the new variable became “None/unknown”, and (3) responses of the rest observations were treated as “Partially prevented”. Then, the responses of the new variable were integrated into dichotomised mating method, creating another new variable of “Implementing BVD control measures on service bulls”. Three responses of the new variables were set as (1) “Bulls not used” if respective mating method was “Bulls not used”, (2) “No” if respective mating method was “Bulls used” but respective BVD prevention status of bulls was not “All prevented”, (3)

“Yes” if respective mating method was “Bulls used” and respective BVD prevention status of bulls was “All prevented”.

Appendix 2

Section 1. A copy of survey administered to New Zealand beef breeding farmers.



BVD FREE NEW ZEALAND
Building grassroots capacity to eradicate bovine viral diarrhoea virus
BVD Management Survey for Beef Herds

Page 1

Dear Farmer,

Bovine viral diarrhoea virus (BVD) is a common infectious disease of cattle that costs New Zealand farmers more than **\$150 million per year** from direct production losses in the **15% of dairy herds** and **65% of beef herds** that are actively infected with the virus. BVD even has significant economic impacts on virus-free herds due to the high ongoing costs of testing and vaccinating cattle to prevent future disease outbreaks.

Several European countries have already launched successful national BVD eradication programmes and we strongly believe that eradicating BVD from New Zealand would be technically feasible and highly profitable for the cattle industries. However, we need a lot more information on how the virus spreads between farms and how much it currently affects herd performance to prepare a sound business case around the different BVD management options (including voluntary control, phased-in mandatory control, and fast-track eradication).

- **Why are we running this survey?**
Every farm in New Zealand has a unique management style and different risk factors for BVD. We specifically want to know how BVD affects your herd and what control measures would be practical for you to put into place. This information will allow us to design and test different national BVD control programmes using an innovative computer simulation model to find the approach that will have the greatest financial benefit at the lowest cost to your farm business.
- **What will the survey involve?**
This BVD Management Survey should take around 30 minutes to complete and has five sections that will ask you questions about your (1) Contact Details and Background, (2) BVD Testing History, (3) Farm Management Practices, (4) BVD Biosecurity Risks, and (5) Opinions Towards National Control. The survey may be completed on paper and either posted or scanned and e-mailed to:

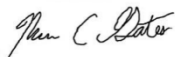
Carolyn Gates (BVD Free New Zealand)
Massey University
Institute of Veterinary, Animal and Biomedical Science (IVABS)
Private Bag 11-222 Palmerston North, New Zealand 4442
Email: c.gates@massey.ac.nz




We will also shortly be offering the option of completing the survey online. Please visit the project website (www.bvdfree.org.nz) for more details.

- **How will we use the information?**
Any information you provide will be treated as strictly confidential and used only for the purpose of understanding how we can better manage BVD in New Zealand. If you have any questions or concerns about participating in the project, please do not hesitate to contact the Project Manager (Carolyn Gates) at c.gates@massey.ac.nz or 06 951 8140.

This Sustainable Farming Fund research project is an exciting opportunity to change how we control infectious diseases in New Zealand cattle industry and we look forward to working with you over the next three years.

Sincerely,





Dr. Carolyn Gates
Project Manager

This research programme was made possible through the generous financial support of Sustainable Farming Fund, AGMARDT, Massey University, MSD Animal Health, and Zoetis. We are grateful to our partners at the National BVD Steering Committee, MPI, DairyNZ, Beef&LambNZ, OSPRI, Gribbles Pathology, NZVP/IDEXX, LIC, SVS Laboratories, ThermoFisher, NZVA, RMPP, and private veterinary practices.



BVD Management Survey for Beef Herds

Page 2

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Section 1: Contact Details and Background Information

Farmer Contact Details

Name	
Street Address	
Town	
Postal code	
E-mail	
Phone	
Veterinary Clinic	
Primary Veterinarian	

Background Information

Are you primarily responsible for making management decisions on the farm?	<input type="checkbox"/> Yes <input type="checkbox"/> No
Gender:	<input type="checkbox"/> Male <input type="checkbox"/> Female <input type="checkbox"/> Prefer Not to Answer
Age:	<input type="checkbox"/> Under 25 <input type="checkbox"/> 25 to 34 <input type="checkbox"/> 35 to 44 <input type="checkbox"/> 45 to 54 <input type="checkbox"/> 55 to 64 <input type="checkbox"/> Over 65 <input type="checkbox"/> Prefer Not to Answer
Number of years farming beef cattle:	_____
Highest education level:	<input type="checkbox"/> High school <input type="checkbox"/> Undergraduate <input type="checkbox"/> Masters <input type="checkbox"/> Doctoral
Ethnic background:	<input type="checkbox"/> Maori <input type="checkbox"/> New Zealand European <input type="checkbox"/> Other: _____ <input type="checkbox"/> Prefer Not to Answer
Are you willing to be contacted to participate in future research studies?	<input type="checkbox"/> Yes <input type="checkbox"/> No



BVD Management Survey for Beef Herds

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Section 2: BVD Testing History

Previous Testing

Do you believe your herd currently has an active BVD infection?		<input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Unsure
Has your herd been screened for BVD within the last 5 years to determine exposure status?		<input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Unknown
If No:	What was the main reason for not performing testing (select ONE)?	<input type="checkbox"/> Unaware of BVD <input type="checkbox"/> Too expensive <input type="checkbox"/> Low perceived impact of BVD <input type="checkbox"/> No intention to control <input type="checkbox"/> Other: _____
If Yes:	How often do you screen for BVD?	<input type="checkbox"/> Annually <input type="checkbox"/> Other: _____
	When was the last screening test?	____ / ____ (mm/yyyy)
	What screening test(s) were used?	<input type="checkbox"/> Sampling 10-15 youngstock to check for antibodies <input type="checkbox"/> Screening all calves in the herd for virus <input type="checkbox"/> Screening all animals in the herd for virus <input type="checkbox"/> Other _____
	What was the result?	<input type="checkbox"/> Negative <input type="checkbox"/> Positive
	If Positive: Was follow up testing performed to identify individual PI animals?	<input type="checkbox"/> Yes <input type="checkbox"/> No
Have you had a known persistently infected animal(s) in the herd within last 5 years?		<input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Unknown (testing not done)
If Yes:	When was the last PI identified?	____ / ____ (mm/yyyy)
	What was the outcome for PI animals?	<input type="checkbox"/> Culled <input type="checkbox"/> Sold <input type="checkbox"/> Remained in herd <input type="checkbox"/> Other: _____



BVD Management Survey for Beef Herds

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Section 3: Farm Management Information

Farm Location

Please provide the NAIT numbers for all locations in **your ownership** where your cattle are grazed

Please provide the NAIT numbers for all locations **NOT in your ownership** where your cattle are grazed-off (i.e. heifer rearers or off-site grazing)

What type of livestock operations are located on your farm? (Check all that apply)

- Beef breeding
- Beef finishing
- Dairy milking
- Sheep
- Dairy grazing
- Deer
- Other _____

Please provide the size of effective grazing area of your enterprise.

_____ hectares

Herd Demographics

Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).

What was the total number of cattle on the farm at the planned start of calving?

Please provide the total number of cattle on the farm by management type at the planned start of calving.

- Breeding cows _____
- Breeding bulls _____
- Store cattle (weaning to slaughter) _____
- Replacement heifers (Rising 1 year olds) _____
- Replacement heifers (Rising 2 year olds) _____
- Other _____



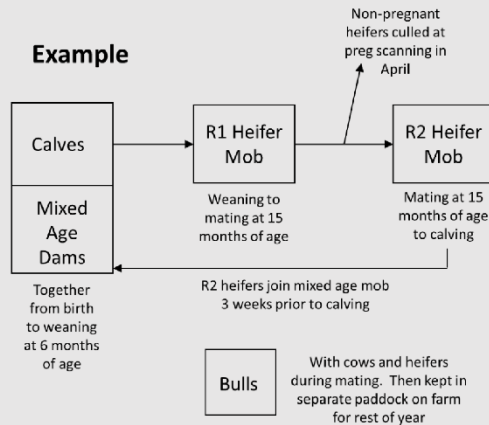
BVD Management Survey for Beef Herds

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Please describe how your herd is managed in different mobs (i.e. which cattle are grazed together at different times of year).

** This is to help us understand how BVD might spread within your herd and what would be the most cost-effective strategy to test your herd for BVD. **

Example





BVD Management Survey for Beef Herds

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What approximate dates during the year are beef cattle yarded for routine management events?

** This is to help us understand when BVD testing and/or vaccination could be integrated into your routine management calendar. **

For Mob, indicate whether they are:

Calves (C), R1 Heifers (R1), R2 Heifers (R2), Mixed Age Cows (MA), or Bulls (B)

Events could include drenching, vaccination, scanning, etc

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____

Date ___/___/___ Mob: ___ Event: _____



BVD Management Survey for Beef Herds

Page 7

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Reproductive Management and Performance

Please answer the following questions related to **calving** during the 2016/17 season (1 July 2016 to 30 June 2017). *This is to help us understand the impact of BVD on herd performance.*

When did calving start in the 2016 season?	_____ / _____ / _____
How many total calves were weaned?	_____
How many total calves were born?	_____
If you know: What was the approximate calving distribution (% of animals calving in each 3 week period)?	Weeks 1 to 3 _____ Weeks 7 to 9 _____ Weeks 4 to 6 _____ Weeks 10 to 12 _____
Were any of the calves stunted or born with birth defects?	<input type="checkbox"/> Yes <input type="checkbox"/> No
If Yes: Please provide additional details around the number of animals and clinical signs	
Did you measure the growth rate of calves from birth to weaning?	<input type="checkbox"/> Yes <input type="checkbox"/> No
If Yes: What was the average daily gain (kg/day)?	_____

Please answer the following questions related to **mating** during the 2016/17 season (1 July 2016 to 30 June 2017).

When did mating start in the 2016 season?	Cows _____ / _____ / _____ Heifers _____ / _____ / _____
How many bulls were used?	_____
How long was the mating period (weeks)?	Cows _____ Heifers _____
How many total females were mated?	Cows _____ Heifers _____
Was pregnancy scanning performed after the end of mating?	<input type="checkbox"/> Yes <input type="checkbox"/> No
If Yes: What was the empty rate (% of animals not pregnant)?	Cows _____ Heifers _____
How many breeding animals were culled?	Cows _____ Heifers _____



BVD Management Survey for Beef Herds

Page 8

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Veterinary Costs and Information Sources

Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).

This is to help us understand the impact of BVD on other animal health issues.

How many cases of scours did you treat?	_____
How much on did you spend for non-routine veterinary care to treat scours and/or respiratory disease? Please exclude the cost of routine vaccination or drenching	NZD _____

What are the top 3 animal health concerns for your beef operation?

1. _____

2. _____

3. _____

On a scale of 1 (no knowledge) to 10 (expert), how would you rate your knowledge of BVD?

Where do you currently receive most of your information about BVD?

- Veterinarian
- Industry Magazines / Publications
- BVD Steering Committee Website
- Other farmers
- Other

What could be done to improve how information about BVD is communicated?



BVD Management Survey for Beef Herds

Visit www.bvdfree.org.nz today to find out more about the research project and how you can get behind BVD control in New Zealand.

Section 4: Biosecurity Information

Purchased Cattle

Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).

Did you purchase any cattle during this time period (including breeding bulls)? Yes No

If Yes: Please provide the approximate number purchased by management type and indicate how many of these were tested for BVD and/or vaccinated for BVD before entering your farm.

	Number purchased	Number of these BVD tested	Number of these BVD vaccinated
Breeding cows			
Breeding bulls			
Calves (pre-weaning)			
Store/finishing cattle (weaning to slaughter)			
Replacement heifers (Rising 1 year olds)			
Replacement heifers (Rising 2 year olds)			

How often were the purchased cattle isolated before being mixed with the herd?

Never Rarely Sometimes Often Always

If you isolate the purchased cattle: How long were the purchased cattle isolated before being mixed? _____ days

How often did you ask about the BVD disease status of the source herd(s)?

Never Rarely Sometimes Often Always

How often did you ask about the BVD vaccination status of the source herd(s)?

Never Rarely Sometimes Often Always

What factors influenced your decisions around BVD biosecurity for purchased cattle?

Empty text box for providing factors influencing BVD biosecurity decisions.



BVD Management Survey for Beef Herds

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Sold Cattle

Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).

Did you sell any cattle to other farms during this time period? Yes No

If Yes: Please provide the number of each type of animal sold

- Breeding cows _____
- Breeding bulls _____
- Calves (pre-weaning) _____
- Store/finishing cattle (weaning to slaughter) _____
- Replacement heifers (Rising 1 year olds) _____
- Replacement heifers (Rising 2 year olds) _____

How often did the buyer(s) ask about the BVD disease status of your herd?

- Never Rarely Sometimes Often Always

How often did the buyer(s) ask about the BVD vaccination status of your herd?

- Never Rarely Sometimes Often Always

Neighbouring Farms

Do you have any neighbour(s) who graze cattle on pastures with shared fenceline boundaries to your cattle? Yes No

If Yes: Please provide the number of properties with cattle sharing boundaries with your farm. _____

Do any of the fenceline boundaries permits direct nose-to-nose contact with neighbouring cattle? Yes No

What is the BVD status of the neighbouring farms?
 All negative
 All positive
 Mixed positive and negative
 Don't know

Do you share yards, pasture, pond or other water sources with your neighbour(s)? Yes No

Do you share equipment, such as dehorners or trailers, with your neighbour(s)? Yes No



BVD Management Survey for Beef Herds

Page 11

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Off-Site Grazing and Movements	
Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).	
Were any of your cattle moved off-site for grazing? <input type="checkbox"/> Yes <input type="checkbox"/> No	
If Yes:	Approximately how long were the animals off-site? _____ (weeks)
Were the animals co-grazed with other herds? <input type="checkbox"/> Yes <input type="checkbox"/> No	
If Yes:	Were the other herds known to be free from BVD? <input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Did not ask
Were any of the animals grazed off-site pregnant at any time during that period? <input type="checkbox"/> Yes <input type="checkbox"/> No	
If Yes:	Were these animals vaccinated prior to conception? <input type="checkbox"/> Yes <input type="checkbox"/> No
	Were the calves from these animals tested for BVD? <input type="checkbox"/> Yes <input type="checkbox"/> No
	Were the animals isolated on their return? <input type="checkbox"/> Yes <input type="checkbox"/> No
Were any cattle from other herds moved onto your site for grazing? <input type="checkbox"/> Yes <input type="checkbox"/> No	
If Yes:	How many herds did the cattle originate from? _____
Were these cattle co-mingled with your stock? <input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Did not ask	
Were these cattle known to be free from BVD? <input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Did not ask	
Were these cattle vaccinated against BVD? <input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Did not ask	
Were other cattle moved off farm and returned for any of the following reasons?	
<input type="checkbox"/> Attending show <input type="checkbox"/> Breeding elsewhere <input type="checkbox"/> Veterinary treatment <input type="checkbox"/> Other _____	



BVD Management Survey for Beef Herds

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BVD Vaccination	
Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).	
Did you vaccinate any of your cattle for BVD? <input type="checkbox"/> Yes <input type="checkbox"/> No	
If No: What was your primary reason for not vaccinating?	<input type="checkbox"/> BVD not present in the herd <input type="checkbox"/> Vaccination too expensive or impractical <input type="checkbox"/> BVD present, but not impacting herd performance <input type="checkbox"/> BVD present, but no intention to control <input type="checkbox"/> Other _____
If Yes: What vaccine product was used?	<input type="checkbox"/> OneShot BVD <input type="checkbox"/> Bovilis <input type="checkbox"/> Bovi-shield Gold <input type="checkbox"/> Other _____
Which groups of animals were vaccinated for BVD?	<input type="checkbox"/> Breeding cows <input type="checkbox"/> Breeding bulls <input type="checkbox"/> Calves (pre-weaning) <input type="checkbox"/> Store/finishing cattle (weaning to slaughter) <input type="checkbox"/> Replacement heifers (Rising 1 year olds) <input type="checkbox"/> Replacement heifers (Rising 2 year olds)
Approximately what dates were the vaccines given?	

Contact with People	
How often do you see uninvited visitors/trampers passing through your property?	
<input type="checkbox"/> Never <input type="checkbox"/> Rarely <input type="checkbox"/> Sometimes <input type="checkbox"/> Often <input type="checkbox"/> Always	
How often do people (e.g. vets, calf dehorners, or scanners) who work with animals wash their boots and other equipment before or after making contact with your cattle?	
<input type="checkbox"/> Never <input type="checkbox"/> Rarely <input type="checkbox"/> Sometimes <input type="checkbox"/> Often <input type="checkbox"/> Always	
How many visits from the following type of personnel occur on average every month?	
	Jan Feb Mar Apr May Jun Jul Aug Sep Oct Nov Dec
Veterinarians	____
Livestock transport vehicles	____

Question continued on next page



BVD Management Survey for Beef Herds

Page 13

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	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
Farm advisors												
Stock agents												

Contact with Sheep

Please answer the following questions related to the 2016/17 season (1 July 2016 to 30 June 2017).

<p>Please provide the average number of each sheep type on your farm.</p>	<p><input type="checkbox"/> No sheep present</p> <p><input type="checkbox"/> Mixed age ewes _____</p> <p><input type="checkbox"/> Two-tooth ewes _____</p> <p><input type="checkbox"/> Ewe lamb/hogget _____</p> <p><input type="checkbox"/> Mixed age rams _____</p>
<p>How many ewes were mated in the 2016 season?</p>	<p>_____</p>
<p>Please provide information on the following production parameters (if available).</p>	<p>Scanning percentage _____</p> <p>Lambing percentage _____</p> <p>Weaning percentage _____</p>
<p>Were any of your cattle directly co-grazed with a sheep flock?</p>	<p><input type="checkbox"/> No</p> <p><input type="checkbox"/> Yes - in the same paddock at the same time</p> <p><input type="checkbox"/> Yes - in adjacent paddocks with possible contact through fenceline</p> <p><input type="checkbox"/> Yes - on the same paddock, but at different times</p>
<p>If Yes: Which management groups were co-grazed with sheep?</p>	<p><input type="checkbox"/> Breeding cows</p> <p><input type="checkbox"/> Breeding bulls</p> <p><input type="checkbox"/> Calves (pre-weaning)</p> <p><input type="checkbox"/> Store/finishing cattle (weaning to slaughter)</p> <p><input type="checkbox"/> Replacement heifers (Rising 1 year olds)</p> <p><input type="checkbox"/> Replacement heifers (Rising 2 year olds)</p>
<p>Has your sheep flock ever been diagnosed with Hairy shaker (Border) disease?</p>	<p><input type="checkbox"/> Yes <input type="checkbox"/> No <input type="checkbox"/> Unsure</p>



BVD Management Survey for Beef Herds

Visit www.bvdfree.org.nz today to find out more about the research project and how you can get behind BVD control in New Zealand.

Section 5: Opinions Towards National Disease Control

Are you aware of the national BVD control programmes that countries in Europe have implemented? Yes No

Do you believe it is possible to eradicate BVD from New Zealand? Yes No Unsure

On a scale of 1 (least supportive) to 10 (most supportive), how strongly do you support having a coordinated national BVD eradication programme in New Zealand? _____

Comments

Which of the following types of coordinated national BVD control programme would you most support?

Voluntary – Decision to control BVD is left entirely to individual farmers

Phased – Level of BVD control is progressively increased from voluntary to compulsory over several years

Compulsory – BVD control is legislated by the government from the start

Comments

Who should decide what approach to national BVD eradication New Zealand should take?

National BVD Steering Committee

Farmers (by vote)

Industry (DairyNZ and Beef&LambNZ)

Other: _____

Comments



BVD Management Survey for Beef Herds

Page 15

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Which of the following BVD control measures would you consider **voluntarily** implementing on your farm?

- Purchasing BVD free or vaccinated animals only
- Isolation of cows brought in to your farm
- Double fencing on your farm boundary
- Testing calves of animals moved off-site during pregnancy
- Testing all replacement calves
- Testing all replacement bulls and heifers
- Only co-grazing cattle with herds that are free from BVD
- Vaccinating at-risk stock against BVD

Comments

Which of the following BVD control measures would you support as part of a **mandated** national BVD eradication programme?

- Screening annually to establish BVD status
- Requiring herds to declare BVD status at the time of sale
- Restricting movements of animals shedding BVD virus
- Establishing a national database to record herd BVD status
- Mandating that BVD positive herds take appropriate measures to control disease

Comments



BVD Management Survey for Beef Herds

Page 16

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What do you see as the biggest **benefits** to eradicating BVD from New Zealand?

What do you see as the biggest **challenges** to eradicating BVD from New Zealand?

What features could be built into an eradication programme to increase success?

Additional Comments

Thank you for your time!

Section 2. Duration of protection against BVDV infection via maternal antibody.

We estimated that the duration of protection against BVDV infection via maternal antibody (Ab) using the reported data of other literature. First, we assumed that the minimum Ab titre required for protection against BVDV infection was 1:8 (Platt et al., 2009; Ridpath et al., 2003). We then extracted the reported data of maternal Ab titre of individual calves over time from two studies (Coria & McClurkin, 1978; Kendrick & Franti, 1974). For the extraction, only the data of animals; (1) with sufficient ($> 1:8$) initial Ab titre, and (2) without any evidence of BVDV infection before the titre dropped to an undetectable level (1:1). Using the extracted data, the decay rate of maternal Ab in the binary logarithm (\log_2) scale was estimated by using generalised linear mixed model with the individual calf id as a random effect. The mean duration of protection against BVDV infection was estimated by calculating the age with maternal Ab titre matching to 3 ($\log_2 8$) based on the estimated decay rate and initial Ab titre (i.e. intercept). To adjust the variation of the titre between calves, we assumed that the duration of protection followed normal distribution, and the standard deviation of the duration was estimated by dividing the standard deviation of the random effect by the coefficient of age.

Table S1. Result of generalised linear mixed model on BVDV maternal antibody titre in binary logarithm scale.

Variable	Coefficient	Standard error	P value
Random effect (Calves)	N/A	1.305 *	< 0.0001 **
Intercept	9.546	0.493	< 0.0001
Age	-0.042	0.002	< 0.0001

Key: N/A, Not available.

* Standard deviation of the random effect

** P value of log-likelihood test

The estimated decay rate of maternal Ab was -0.042 which was in accordance with previously estimated values in other studies (Table S1) (Downey et al., 2013; Zimmerman

et al., 2006). The estimated initial maternal Ab titre was 9.546 (in the binary logarithm scale), and the mean duration of protection against BVDV infection was 154.8 days. The standard deviation of the random effect was 1.305, and the standard deviation of the duration of protection was 30.8 days. Therefore, the mean and standard deviation of the duration of protection against BVDV infection via maternal Ab was approximated to 155 and 31 days, respectively (Figure S1).

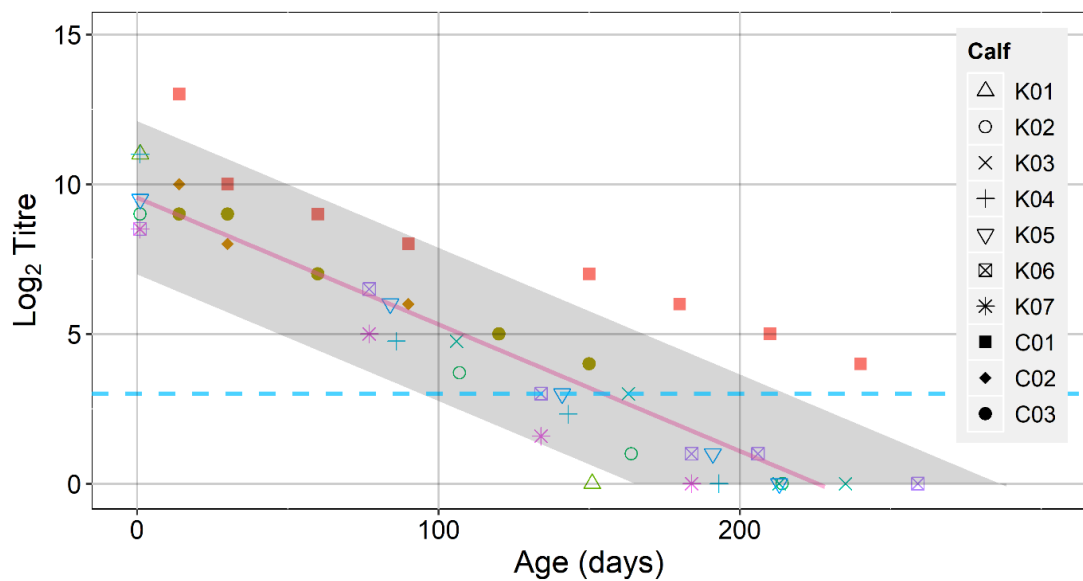


Figure S1. Decay rate of BVDV maternal antibody titre in binary logarithm scale. Blue dashed horizontal line indicates the minimum antibody titre required for protection against BVDV infection.

Section 3. Detailed algorithm of approximate Bayesian computation-sequential Monte Carlo (ABC-SMC).

To estimate unknown parameters for each of 9 beef herds, we implemented ABC-SMC algorithm in the following order. Since we estimated multiple parameters at the same time, θ below is the vector of parameters.

- (1) Set $t = 1$, where t is the SMC sequence indicator. Initialise the threshold values for the first round ($t = 1$), $\varepsilon_1(1)$ and $\varepsilon_2(1)$, by;

- a. Running the simulation model 2,000 times without rejecting any particles, and estimating the distances (D_k) of summary statistics as;

$$D_k = \sqrt{\sum_{i=1}^n (T_i(+)^k_{obs} - T_i(+)^k_{sim})^2}$$

where n is the number of herds, $T_i(+)^1_{obs}$ and $T_i(+)^1_{sim}$ are the observed and simulated number of test positive heifers in the first sampling round for herd i , respectively, $T_i(+)^2_{obs}$ and $T_i(+)^2_{sim}$ are the observed and simulated number of seroconverted heifers in the second sampling round for herd i , respectively.

- b. Setting $\varepsilon_1(1)$ and $\varepsilon_2(1)$ as the median values of D_1 and D_2 , respectively.

(2) Set $i = 1$, where i is the particle indicator.

(3) Generate a particle of parameter set, θ , by;

- a. If $t = 1$, sample θ^{**} from $\pi(\theta)$, where $\pi(\theta)$ is the prior distributions of θ .
- b. If $t > 1$, sample θ^* from the particles of previous sequence, $\{\theta_{t-1}\}$, with weights, $\{w_{t-1}\}$. Then perturb the particle $\theta^{**} \sim K(\theta|\theta^*)$, where $K(\cdot)$ is a perturbation kernel. In this study, we used a component-wise Gaussian kernel with the variance as 0.68 times of the variance of particles in the previous SMC sequence. If the probability of $\pi(\theta^{**})$ equals 0, return to (3).

(4) Run the simulation model with the generated particles and calculated D_k .

(5) Accept θ^{**} as the particle θ_t^i if $D_1 < \varepsilon_1(t)$ and $D_2 < \varepsilon_2(t)$, otherwise return to (3).

(6) Calculate weight for the particle, w_t^i , as;

- a. If $t = 1$, $w_t^i = 1$,
- b. If $t > 1$, $w_t^i = \pi(\theta_t^i) / \sum_{j=1}^N w_{t-1}^j K(\theta_t^i, \theta_{t-1}^j)$.

(7) Set $i = i + 1$, and repeat (3) ~ (6) until $i = 2,004$. It indicates that 2,000 particles for each parameter were accepted for one SMC sequence.

(8) Normalise the weights (divide the weight by the sum of weight) and calculate the effective sample size (ESS) as;

$$ESS = \frac{1}{\sum_1^{2000} (\bar{w}^i)^2}$$

where \bar{w}^i was normalised weight.

(9) Calculate new threshold for the next sequence, $\varepsilon_1(t + 1)$ and $\varepsilon_2(t + 1)$ from the median values of D_1 and D_2 , respectively, and set $t = t + 1$. Return to (1) until $t = 15$.

In total, we sampled approximately 97.2 million particles and ESS of any parameter during the whole sequence was between 795 and 2,000.

Section 4. General sensitivity analysis of model behaviour.

Since the simulation model in this study was newly developed, we investigated the model behaviour with respect to different parameter values. The sensitivity analysis was conducted using a complete factorial design on five unknown parameters; (1) within-herd BVDV transmission rate for PI animals (β_p), (2) the initial proportion of BVDV seropositive animals (μ), (3) the proportion of introduced PI animals (ρ), (4) the day of PI animals being introduced (τ), and (5) the proportional difference of β_T compared to β_p . Five values (0.1, 0.3, 0.5, 0.7, and 0.9) were tested for all parameters except τ , and the days equivalent to 10th, 30th, 50th, 70th, and 90th percentile of the duration between the day of weaning (day 0) and the day of first sampling were used for τ , resulting in testing 3,125 different scenarios. Each scenario was simulated for 1,000 times to adjust the variation caused by stochasticity. For each simulation, we stored the parameter values used and the number of seropositive heifers at the first and second sampling events. Once whole simulation was over, a dataset of 3,125,000 observations was created, and the numbers of seropositive heifers were separately analysed using multivariable Poisson regression with the parameters as explanatory variables. For the simplicity, the analysis was conducted under the management features of Farm 1 only.

Table S2. Multivariable Poisson regression on the number of seropositive heifers at two sampling events.

Variable	Number of seropositive heifers at;			
	First sampling		Second sampling	
	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value
Intercept	0.688 (0.687, 0.689)	< 0.0001	0.764 (0.763, 0.765)	< 0.0001
β_P	1.050 (1.049, 1.052)	< 0.0001	1.025 (1.024, 1.026)	< 0.0001
μ	1.756 (1.753, 1.758)	< 0.0001	1.525 (1.523, 1.526)	< 0.0001
ρ	0.595 (0.594, 0.595)	< 0.0001	0.659 (0.658, 0.660)	< 0.0001
τ	1.000 (1.000, 1.000)	< 0.0001	1.000 (1.000, 1.000)	< 0.0001
K	1.014 (1.013, 1.015)	< 0.0001	1.005 (1.004, 1.006)	< 0.0001

Key: CI, confidence interval; K, the proportional difference of β_T compared to β_P .

Table S2 illustrates the results of multivariable Poisson regression models. Not surprisingly, the number of seropositive heifers at both the first and second sampling were the most sensitive to μ , followed by ρ . Compared to μ or ρ , β_P, τ and the proportional difference of β_T compared to β_P had only a limited impact on the number of seropositive heifers in both sampling occasions. This result indicates that the posterior distribution of β_P using ABC-SMC in this study is expected to be mainly affected by the values of μ and ρ since the estimation of β_P in ABC-SMC of this study depended on the number of seropositive heifers at the first sampling and the number of seroconverted heifers at the second sampling events.

Section 5. Number of sampled and test-positive heifers at each sampling, herd size, breeding period, and day of both sampling events for each 9 New Zealand beef breeding farms.

Table S3. Number of sampled and test-positive heifers at each sampling, herd size, breeding period, and day of both sampling events for each 9 New Zealand beef breeding farms.

Farm	Herd size	# Sampled	# 1 st positive	# 2 nd positive	Breeding period †	Day of 1 st sampling †	Day of 2 nd sampling †
1	40	15	9	6	250 ~ 306	590	735
2	90	14	11	1	230 ~ 265	549	690
3	66	15	2	6	203 ~ 245	533	711
4	23	15	13	2	236 ~ 292	579	731
5	24	15	13	2	571 ~ 641	946	1,097
6	37	14 *	7	4	250 ~ 292	625	752
7	20	15	3	9	590 ~ 646	950	1,097
8	20	15	12	3	236 ~ 320	597	736
9	29	15	14	1	240 ~ 261	640	794

* Numbers of sampled replacement heifers at the first round adjusting for the censored heifers at the second sampling round. Originally 15 heifers were sampled at the first round.

† Day of each event was measured from day 0 which was the day of heifers being weaned.

Section 6. Evolution of parameter values and Inspection of estimated parameters.

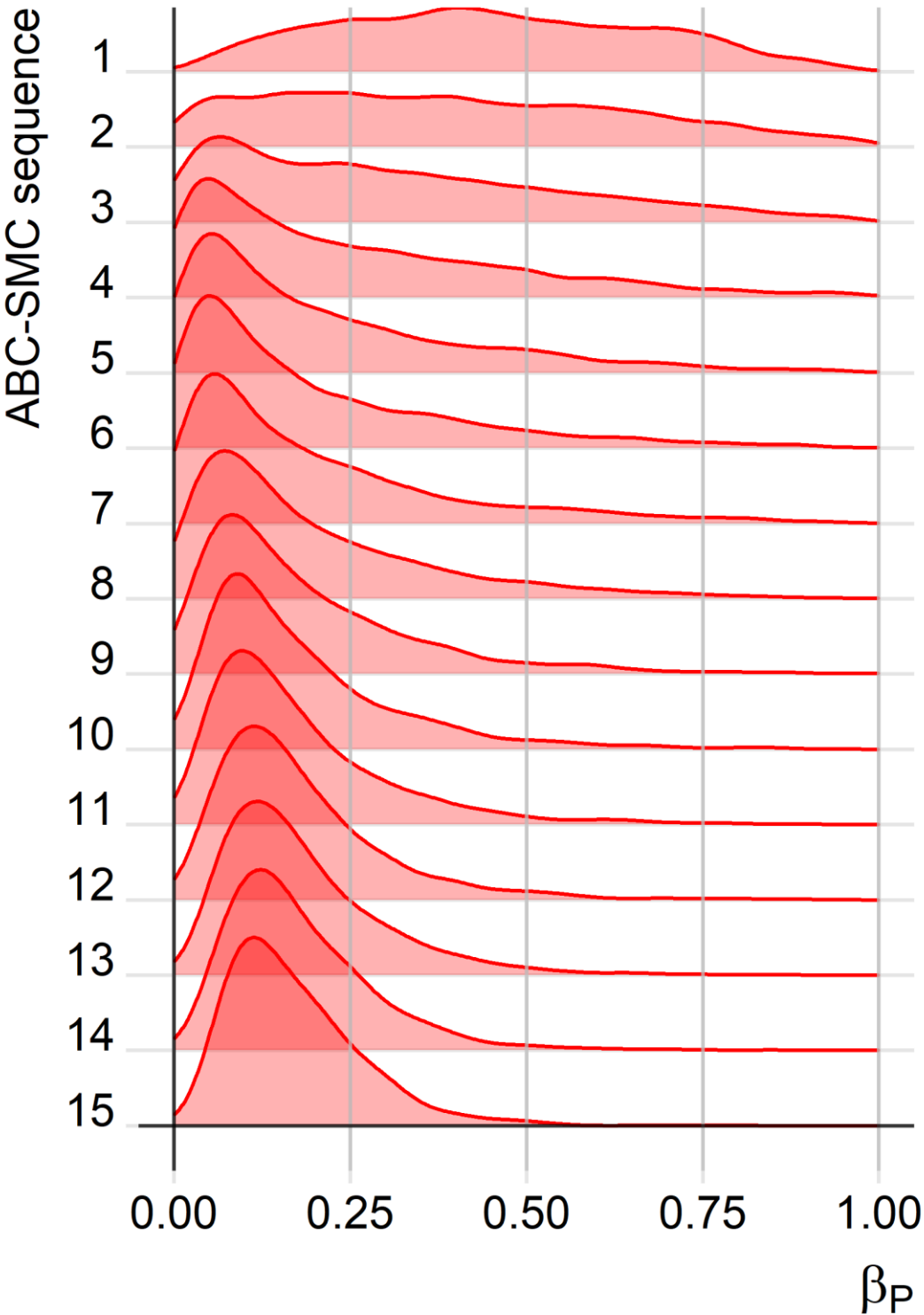


Figure S2. Evolution of β_P values over 15 ABC-SMC sequences. The mode (95% highest posterior density range) of β_P in the final sequence was 0.11 (0.03 ~ 0.34).

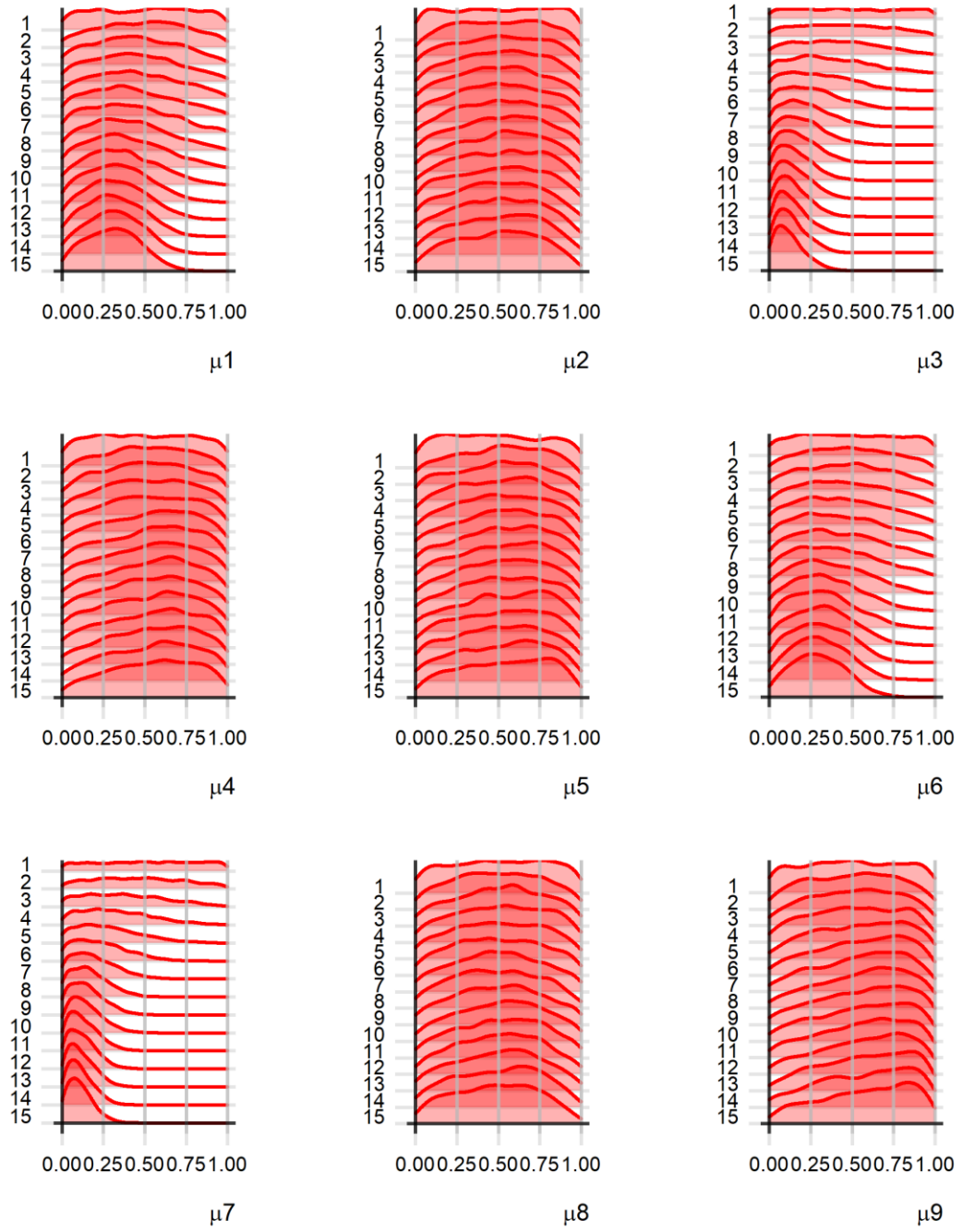


Figure S3. Evolution of μ values for 9 studied farms over 15 ABC-SMC sequences.

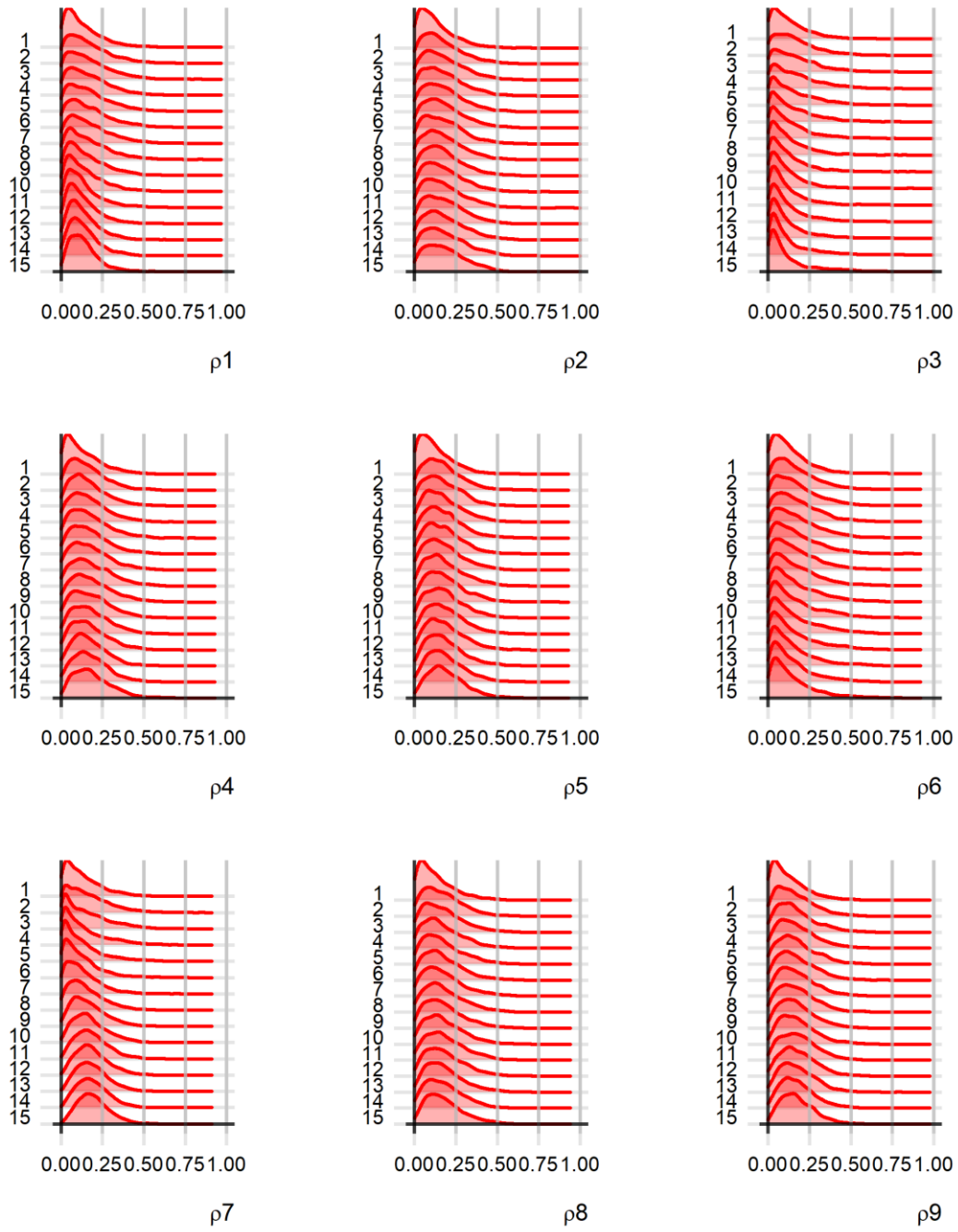


Figure S4. Evolution of ρ values for 9 studied farms over 15 ABC-SMC sequences.

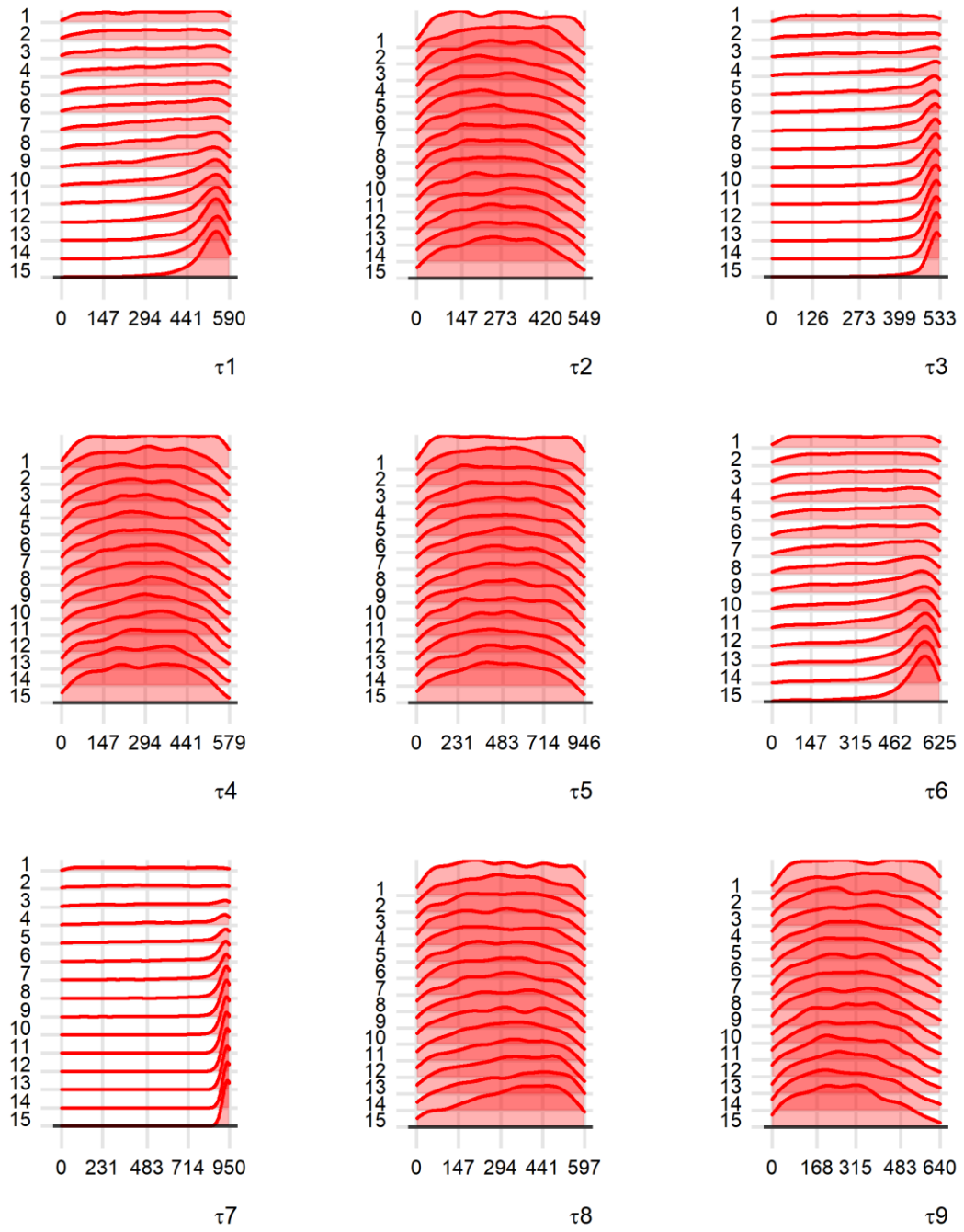


Figure S5. Evolution of τ values for 9 studied farms over 15 ABC-SMC sequences.

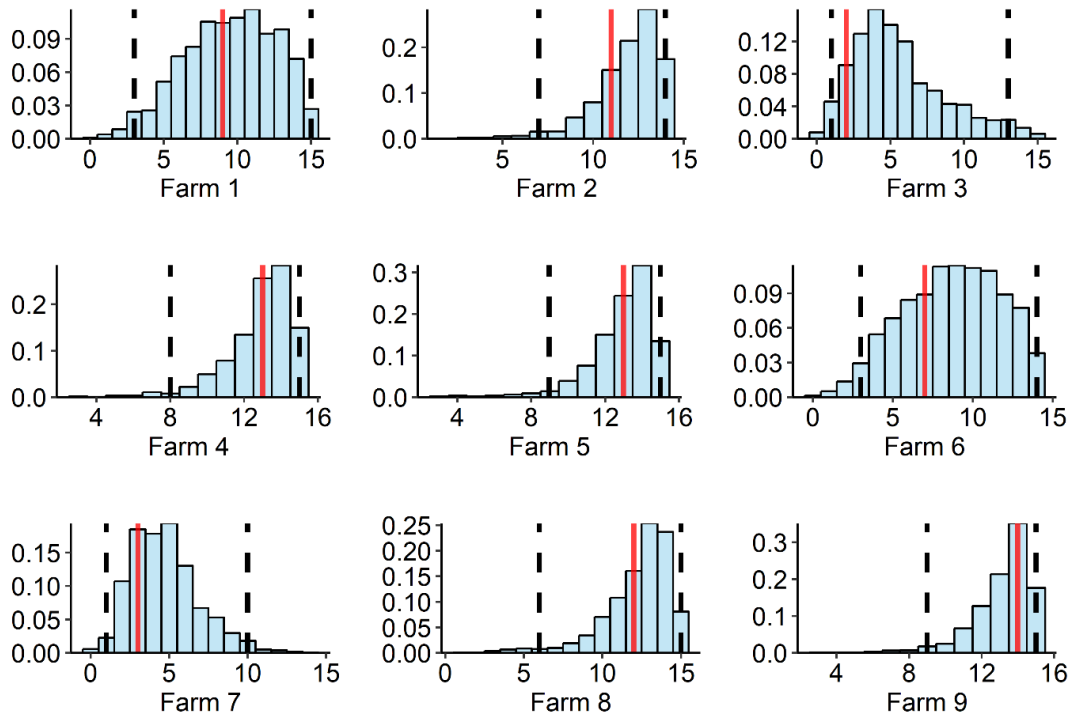


Figure S6. Observed (red line) and simulated (blue histogram) number of test positive heifers in the first sampling round. The distribution of simulated summary statistics was drawn by applying 2,000 randomly sampled parameter values from the relevant posterior distributions. Black dashed lines indicate the 95% prediction interval of the simulated distribution.

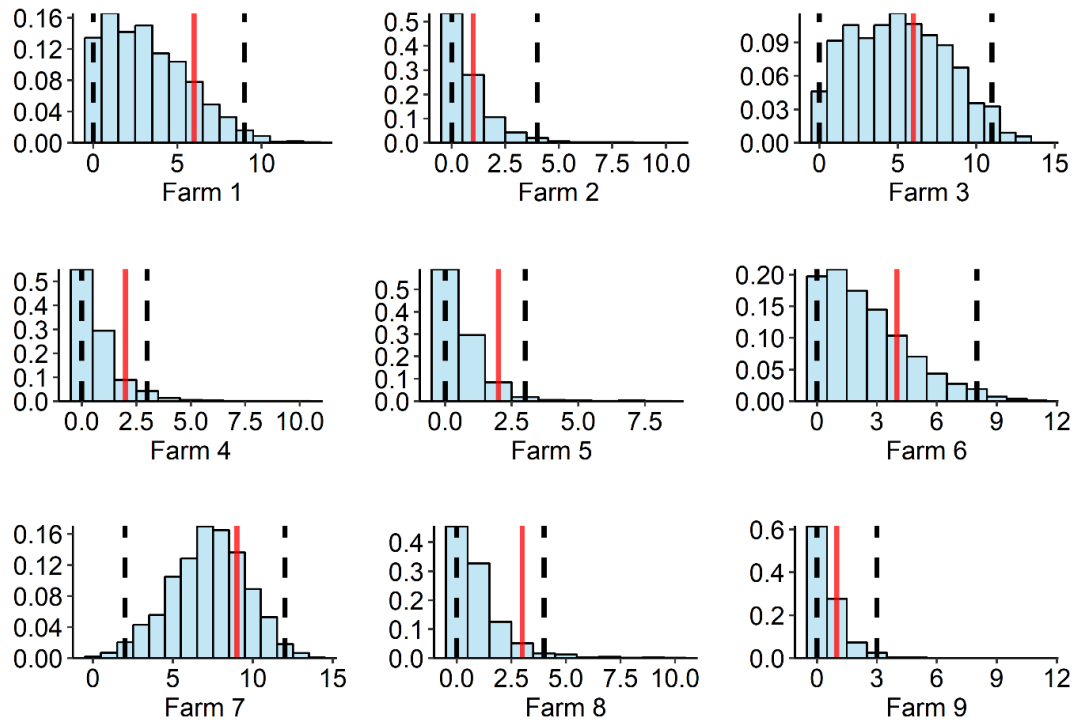


Figure S7. Observed (red line) and simulated (blue histogram) number of seroconverted heifers in the second sampling round. The distribution of simulated summary statistics was drawn by applying 2,000 randomly sampled parameter values from the relevant posterior distributions. Black dashed lines indicate the 95% prediction interval of the simulated distribution.

Section 7. Sensitivity analysis.

We investigate the impact of using a uniform prior of the proportion of introduced PI animals ($\pi(\rho)$) in this study. Given the purpose was to measure the impact of uniform prior, we collected 1,000 particles for each parameter for the sake of computational efficiency.

Compared to using an informative $\pi(\rho)$, some but not all parameters showed different posterior distribution (Table S4). Especially, the posterior distributions of ρ were overestimated in general compared to the original study. This result of sensitivity analysis would indicate that our model was not identifiable with current information (i.e. data and priors). We believe that the identifiability issue was due to the lack of enough observation — the number of tests as well as the sample size of each test — per farm, which could be warranted in the future study.

Table S4. Mode (95% highest posterior density region) of estimated parameters with a uniform prior of the proportion of introduced PI animals. Rightmost column shows the posterior values of parameters in the original study.

Parameters	$\pi(\rho) = U(0,1)$	$\pi(\rho) = B(1.14,7.79)$
β_P	0.21 (0.03, 0.75)	0.11 (0.03, 0.34)
μ_1	0.42 (0.05, 0.66)	0.29 (0.01, 0.59)
μ_2	0.54 (0.12, 0.94)	0.54 (0.05, 0.91)
μ_3	0.04 (0.00, 0.24)	0.06 (0.00, 0.28)
μ_4	0.72 (0.19, 0.98)	0.51 (0.09, 0.99)
μ_5	0.72 (0.15, 1.00)	0.54 (0.03, 0.93)
μ_6	0.22 (0.00, 0.53)	0.32 (0.00, 0.58)
μ_7	0.03 (0.00, 0.25)	0.08 (0.00, 0.24)
μ_8	0.59 (0.11, 0.94)	0.55 (0.00, 0.87)
μ_9	0.82 (0.26, 1.00)	0.82 (0.09, 0.99)
ρ_1	0.14 (0.02, 0.89)	0.08 (0.01, 0.30)
ρ_2	0.49 (0.02, 0.89)	0.09 (0.00, 0.43)
ρ_3	0.93 (0.09, 1.00)	0.03 (0.00, 0.38)
ρ_4	0.25 (0.04, 0.88)	0.15 (0.00, 0.39)
ρ_5	0.27 (0.01, 0.83)	0.12 (0.02, 0.40)
ρ_6	0.65 (0.78, 1.00)	0.04 (0.00, 0.39)
ρ_7	0.24 (0.02, 0.59)	0.17 (0.03, 0.38)
ρ_8	0.31 (0.02, 0.89)	0.11 (0.00, 0.39)
ρ_9	0.23 (0.01, 0.86)	0.11 (0.00, 0.37)
τ_1	567 (109, 590)	546 (377, 590)
τ_2	378 (33, 525)	378 (19, 506)
τ_3	525 (161, 533)	525 (406, 533)
τ_4	315 (10, 525)	315 (18, 522)
τ_5	483 (28, 867)	378 (46, 880)
τ_6	609 (45, 622)	567 (202, 625)
τ_7	945 (925, 950)	945 (912, 950)
τ_8	420 (27, 578)	462 (50, 581)
τ_9	378 (13, 567)	315 (6, 540)

Key: $\pi(\cdot)$, Prior distribution; B, beta; U, uniform; ρ , the proportion of introduced PI animals

Appendix 3

Section 1. Estimation of demographic parameters

1.1. Probabilities of conception in normal, TI, and PI animals

1.1.1. Material and method

We estimated the probabilities of conception by artificial insemination (AI) and by mating with bulls by using a stochastic individual-based model of simulating reproduction of 100 dairy cows. In the simulation, the oestrus cycle of each cow was randomly selected between 18 and 24 days for each cycle, and animals either calved or had abortion were assumed to have their first “detectable” oestrus after 15 ~ 49 days plus any days between 18 and 24 days. Gestation period was 281 days, and a natural abortion rate of 0.0001 per day during the whole gestation period was applied to achieve 3.5% of abortion probability that not potentially due to any disease in New Zealand dairy farms (Weston et al., 2012). Each model was iterated 1,000 times to adjust the variability caused by stochasticity.

Since the dairy cows were artificially inseminated for the first 6 weeks, we first estimated the probability of conception by AI. Initially, we assumed all cows were empty and in oestrus at the planned start day of calving (day 0). From day 84 to day 125 (6 weeks), animals in oestrus were detected with 78.6% probability (Anonymous 2018b), and only detected animals were conceived with $x\%$ of probability where x was manually searched. We ran the simulation for 8 years, where the first 3 year was treated as a burn-in period to minimise the effect of initial condition. We compared the mean proportion of cows calved for each calving period during the last 5 years of simulation to the industry-reported value (66.1% of cows bred being pregnant by 6 weeks) (Anonymous 2018b), and repeated the process with varying $x\%$ until the mean proportion approximately matched to the 66.1%.

Once the probability of conception by AI was estimated, the probability of conception by mating with bulls for the next 6 weeks was also estimated using the same method. After the period of AI, animals in oestrus were conceived with $y\%$ for the next 6 weeks (breeding

period with bulls). We repeated the process until the mean proportion of cows calved for each calving period during the last 5 years with $y\%$ of conception probability being matched to approximately 15.0% of cows bred being empty by the calving period (Anonymous 2018b).

We applied the same method to 100 beef cows to estimate the probability of conception by mating with bulls with the same natural abortion rate (Hickson et al., 2012). The only differences were; (1) breeding period started 81 days after the planned start date of calving, (2) breeding period was from day 81 to day 143 (9 weeks), and (3) the probability of conception was searched to match approximately to 91.0% of cows bred being pregnant by pregnancy scanning which was on day 186 (McFadden et al., (2015).

To estimate the risk of conception in transiently infected (TI) animals, we analysed the reported data from two experimental studies that challenged BVDV before/after breeding (McGowan et al., 1993; Whitmore et al., 1981). Using the data, we calculated the risk ratio of conception of each test group, and estimated the weighted-mean using the number of samples and controls for each ratio. For the conception of persistently infected (PI) animals, we assumed that the risk ratio of conception in PI animals is equivalent to the one in TI animals.

1.1.2 Results

The probability of conception by AI in dairy cows was estimated as 60.0%. The estimated probability of conception by mating with bulls was 45.0%. In beef cows, the estimated probability of conception by mating with bulls was 61.0%.

Whitmore et al., (1981) challenged BVDV to 3 groups of cows after 35 days from the day of breeding, and observed 60.0% (9/15), 27.0% (4/15), and 67.0% (10/15) of conception probability depending on the challenge method, while reporting 67.0% (10/15) of conception probability in the control group. The risk ratio of conception in these groups

was 0.896, 0.403, and 1.000, respectively. In another study, McGowan et al., (1993) challenged BVDV to 2 groups at near breeding (before 9 days or after 4 days from the day of breeding), and reported 60.0% (9/15) and 44.4% (8/18) of conception probability, while 78.6% (11/14) of animals in the control group were conceived. The risk ratio of these groups was 0.764 and 0.566, respectively. Using the number of animals in each group and control, the weighted-risk ratio of conception in TI animals was 0.681.

1.2. Daily rate of abortion of normal, TI, and PI cattle

According to Weston et al., (2012), 79 out of 2,246 dairy cows and first-calving heifers had abortion without any relevant clinical findings. Therefore, we estimated the probability of natural abortion as 3.5%. For the simplicity of modelling, we assumed that the abortion can occur during the whole gestation period (281 days), and the daily rate of natural abortion (μ_n) was;

$$\mu_n = -\frac{\ln(1 - 0.035)}{281} = 0.00013/\text{cattle}/\text{day}$$

In this study, we assumed that the probability of natural abortion in beef farms was the same as the one in dairy farms.

For BVDV-induced abortion in TI animals, abortion was assumed to occur at the day of infection for the sake of simplicity. McGowan et al., (1993) challenged BVDV to 2 groups of cows at near breeding (before 9 days or after 4 days from the day of breeding), and reported that 6.7% (1/15 after adjusting 9 conception failure) and 16.7% (3/18 after adjusting 8 conception failure) of cows had abortion, while the control group showed no abortion (0/14 after adjusting 11 conception failure). Using the number of samples per group and the number of controls, the weighted-mean of the probability of abortion if BVDV-infection occurred during 0 ~ 41 days of gestation was estimated as 12.1%. To estimate the probability of abortion when animals were infected with BVDV during the

mid-gestation period (42 ~ 150 days), we used the same method to the reported data of 5 studies (Grooms et al., 2007; McClurkin et al., 1984; Viet et al., 2004; Walz et al., 2017, 2018). Also, we treated calves with congenital defect as abortion for the convenience of the simulation by assuming that calves with congenital defects died when they were born. The reported probabilities of abortion in the previous studies were; 20.0% (3/15) for McClurkin et al., (1984), 6.7% (1/15) for Grooms et al., (2007), 13.3% (2/15) for (Walz et al., 2017), and 12.5% (2/16) for (Walz et al., 2018). The weighted-mean probability was 11.9%. On the estimated probability of abortion, we added 6.3% which was the probability of calves (if dams were infected during their mid-gestation but did not have abortion) expected to be born with congenital defects based on (Viet et al., 2004). Therefore, the overall probability of abortion if BVDV-infection occurred during 42 ~ 150 days of gestation was 17.4%. Finally, we assumed that no abortion occurred when animals were BVDV infected after 150 days of gestation.

It has been speculated that BVDV induces mild placental lesions in dams which causes damage to fetus and eventually results in abortion (Baszler et al., 1995; Liebler-Tenorio, 2005). Based on this speculation, abortion can occur any day of early-mid pregnancy as long as BVDV persists in uterine tissue, which makes plausible that the probability of abortion in PI animals be higher than in TI animals since the virus only exists temporarily in TI animals. Therefore, in order to parameterise the abortion of PI animals, the probabilities of abortion in TI animals were converted into daily abortion rates using the average days of infection (15 days). The converted rates were applied to PI animals during the gestation period (0~41 days: 0.0086/cattle/day, 42~150 days: 0.0128/cattle/day), which was equivalent to 92.5% of abortion probability until 150 days of gestation (excluding the natural abortion).

1.3. Daily rate of natural mortality of calves

Cuttance et al., (2017) reported that approximately 9.5% calves born in a dairy farm naturally died before weaning. Assuming that the weaning occurs at 70 days old, the daily natural mortality rate of dairy calves (m_d) was estimated as;

$$m_d = -\frac{\ln(1 - 0.095)}{70} = 0.0014/\text{calf/day}.$$

In New Zealand beef farms, C. A. Morris et al., (1986) reported that the probability of calf mortality differed by the age of dam. For heifers, the study showed that 14.3% calves died by weaning. Using the same formula above while assuming calves were weaned by 180 days, the daily natural mortality rate of beef calves from heifer-dams (m_h) was equivalent to 0.0009/calf/day. This value can be validated by comparing to a more recent study which reported that the number of survived calves at marking (generally after 2~4 months from the planned start date of calving) being approximately 77.0% of the number heifers bred (Hickson et al., 2012). The study estimated that the proportion of heifers bred being pregnant as 83.0%, and using the estimated proportion, the probability of calf mortality by marking was inferred as 7.2% ($1 - 0.77/0.83$). When we calculated the same value using our estimated daily mortality rate of heifer-born calves (while assuming the marking occurs by 90 days), the probability of calf mortality by marking was 7.4% ($1 - e^{-0.0009 \times 90}$), which was close to the inferred value. For the calves born from beef cows, C. A. Morris et al., (1986) reported that the probability of natural mortality by weaning was 5.0%, which was converted as 0.0003/calf/day (m_m).

Section 2. Estimation of the live-weight of New Zealand cattle.

We calculated the live-weight of each individual cattle by estimating the live-weight gain per day (LWG/day) of each animal for every day and adding it to their weight of previous day. The LWG/day was a function of animal's age, and a formula to estimate LWG/day was calculated by fitting the average birth, weaning, and maximum weights of

cattle in each production type into the Von Bertalanffy function (Brown et al., 1976; García-Muñiz et al., 1998).

According to the Von Bertalanffy function, the weight of cattle in day t (W_t) is defined as;

$$W_t = A \cdot (1 - B \cdot e^{-K \cdot t})^3$$

where A is asymptotic or population-average maximum weight, and B and K are the parameters to be estimated. By differentiating the formula above by t , weight gain or growth rate per t can be calculated as;

$$3 \cdot K \cdot t \cdot \left(\frac{B \cdot e^{-K \cdot t}}{1 - B \cdot e^{-K \cdot t}} \right).$$

For dairy cattle, the birth weight was assumed to be 36.1kg (Hickson et al., 2015) with asymptotic maximum weight of 540kg (Anonymous, 2018a). With these assumptions, B was estimated as 0.5941. By further assuming that the dairy cattle reach 90kg by 70 days old, we calculated K as 0.0040. Therefore, the formula to estimate LWG/day for dairy cattle was;

$$0.0119 \cdot Age \cdot \left(\frac{0.5941 \cdot e^{-0.0040 \cdot Age}}{1 - 0.5941 \cdot e^{-0.0040 \cdot Age}} \right).$$

To add variability of live-weight between animals, we assumed that the birth weight of the calves was normally distributed. For female dairy calves, the mean and standard deviation of birth weight were 38.4 and 0.5kg, respectively (Hickson et al., 2015). For the male calves, the counterparts were 41.8 and 0.5kg, respectively (Hickson et al., 2015).

The same method was applied to beef cattle with 39.2kg of birth weight (S. Morris et al., 2006), 570kg of asymptotic maximum weight, and 228kg of weaning weight (180 days old) (Geenty & Morris, 2017). With these assumptions, the formula to estimate LWG/day for beef cattle was;

$$0.0135 \cdot Age \cdot \left(\frac{0.5903 \cdot e^{-0.0045 \cdot Age}}{1 - 0.5903 \cdot e^{-0.0045 \cdot Age}} \right).$$

Birth weight of female beef calf was randomly selected from a normal distribution with the mean and standard deviation of 39.8 and 0.6, respectively (Hickson et al., 2015). The counterparts of male beef calf were 43.2 and 0.6, respectively (Hickson et al., 2015).

In both dairy and beef models, the weight gain of animals was assumed to stop once they reached relevant maximum weights.

Section 3. Estimation of the efficacy of BVDV killed vaccine.

According to Ridpath, (2013), a higher level of immune response is required to protect fetal infection against BVDV than the one required to prevent clinical diseases. Therefore, we assumed that second dose of BVDV killed-vaccine can be resulted in three outcomes; (1) vaccination conferred not enough immunity to prevent clinical diseases (V_F , vaccine failure), (2) vaccination conferred enough immunity to prevent clinical diseases but to prevent fetal infection insufficiently (V_I , insufficient fetal protection), and (3) vaccination conferred enough immunity to prevent fetal infection (V_S , sufficient fetal protection). We assumed that 7.2% of animals that vaccinated with a killed-vaccine acquire enough immunity to prevent fetal infection ($P(V_S)$) based on an assumption that vaccinated animals showing no signs of the disease after being exposed to BVDV was the evidence of enough immunity to prevent fetal infection (Downey-Slinker et al., 2016). We also assumed that 19.3% of vaccinated animals have vaccine failure ($P(V_F)$), based on an assumption that the presence of both lymphocytopenia and thrombocytopenia indicates not enough immunity to prevent clinical diseases (Downey-Slinker et al., 2016). Therefore, 73.4% of vaccinated animals were assumed to have immunity with insufficient fetal protection ($P(V_I)$).

We then parameterised the impact of BVDV infection on the reproduction of V_I animals based on the reported data of other literature. In a recent meta-analysis study, Newcomer et al., (2015) reported that BVDV vaccination using killed-vaccine reduces the risk of

abortion by 0.662 times. Since BVDV infection during 0 ~ 41 days of gestation was assumed to cause abortion with 12.1% of probability in this study (Section 1 of Appendix 3), we calculated the overall probability of abortion if an animal was infected during 0 ~ 41 days of gestation given the animal was vaccinated ($P_E(Abort|Vacc)$) as 8.0% (0.121×0.662). Assuming that BVDV-induced abortion only occurred in V_F and V_I , $P_E(Abort|Vacc)$ was described as;

$$P_E(Abort|Vacc) = P_E(Abort|V_F) \cdot P(V_F) + P_E(Abort|V_I) \cdot P(V_I) = 0.080$$

where $P_E(Abort|V_X)$ was the probability of abortion in vaccinated animals with the vaccination outcome of x (either complete failure or insufficient protection), and $P(V_F)$ and $P(V_I)$ were the probabilities of vaccine failure (0.193) and insufficient fetal protection (0.734), respectively, among vaccinated animals. Since the animals with vaccine failure are no different to susceptible animals with respect to BVDV infection, we assumed that $P_E(Abort|V_F)$ was 12.1%, therefore, $P_E(Abort|V_I)$ was estimated as 7.7%.

For BVDV infection during 42 ~ 150 days of gestation, the probability of abortion was estimated as 11.9% (Section 1 of Appendix 3), and the overall probability of abortion if an animal was infected during 42 ~ 150 days of gestation given the animal was vaccinated ($P_M(Abort|Vacc)$) was calculated as 7.9% (0.119×0.662). Using the same approach as above, $P_M(Abort|Vacc)$ was then;

$$P_M(Abort|Vacc) = P_M(Abort|V_F) \cdot P(V_F) + P_M(Abort|V_I) \cdot P(V_I) = 0.079$$

, and $P_M(Abort|V_I)$ was estimated as 7.6%.

We also assumed that PI calves can be born if dams were V_I . (Newcomer et al., 2015) reported that the risk of fetal infection in animals inoculated with killed-vaccine was 0.236 times lower than the one without vaccination. Since PI animals can be generated if dams were infected during 42 ~ 150 days but did not have abortion, we described the risk ratio of fetal infection in the vaccinated animals as;

$$P(PI|No\ abort, Vacc) \times P(No\ abort, Vacc)/P(PI|No\ abort, No\ vacc) \times \\ P(No\ abort, No\ vacc).$$

Assuming 87.5% of calves born from dams that infected with BVDV during 42 ~ 150 days of gestation become PI animals (Viet et al., 2004), the denominator of the formula above was $0.875 \times (1 - 0.119) = 0.771$. Also, $P(No\ abort, Vacc)$ in the numerator can be expressed as;

$$P_M(No\ abort|V_F) \cdot P(V_F) + P_M(No\ abort|V_I) \cdot P(V_I) \\ = (1 - 0.119) \cdot 0.193 + (1 - 0.076) \cdot 0.734 = 0.849.$$

Therefore, $P(PI|No\ abort, Vacc)$ was 21.4% ($= 0.236 \times 0.771 \div 0.849$).

According to Grooms et al., (2007), 3 out of 11 non-PI new born calves from vaccinated dams that were BVDV-infected during 52 ~ 150 days of gestation had detectable level of BVDV antibody titre. Therefore, we assumed that 27.3% (3/11) of new born calves from V_I that infected during 42 ~ 150 days of gestation were already BVDV-recovered when they were born (if they were not PI animals), and the rest of calves were assumed to have passive immunity via maternal antibody. The probability of congenital defect of calf from V_I animals was ignored as it has been rarely reported from other studies (Walz et al., 2018).

With our parameterisation, the overall probability of vaccinated dams to calve a PI animal was 30.6%, which was similar to the reported BVDV vaccine efficacy that commercially available in New Zealand (McArthur, 2004; Packianathan et al., 2017).

Section 4. Sensitivity analysis.

Table S5. The impact of parameters on model outputs (number of transiently infected and persistently infected animals, and total gross farm revenue) and contribution of each parameter to output variation.

Parameters	Number of TIs	Number of PIs	Total GFR
Dairy farm	Coef (SE)	Coef (SE)	Coef (SE)
	RCV	RCV	RCV
Within-herd BVDV transmission rate by a PI animal (β_P)	0.92 (0.00)	0.07 (0.00)	-30,100.2 (331.0)
	43.4%	24.6%	9.4%
Relative proportion of β_B to β_P	0.39 (0.00)	0.04 (0.00)	-18,476.7 (418.7)
	4.8%	4.7%	2.2%
Relative proportion of β_F to β_P	0.22 (0.00)	0.02 (0.00)	-5,807.0 (418.7)
	1.5%	1.2%	0.2%
Prevalence of PI animals in the neighbouring farms (K_F)	3.14 (0.03)	0.26 (0.00)	-100,266.3 (4,187.0)
	3.1%	2.5%	0.6%
Mortality of PI animals (μ)	-180.25 (0.69)	-19.95 (0.09)	11,460,929.9 (84,227.2)
	25.6%	35.3%	21.1%
Number of introduced Trojan cows as a relative proportion to T_M (ρ)	6.51 (0.03)	0.74 (0.00)	-798,682.4 (3,310.1)
	21.6%	31.7%	66.4%
Beef farm	Coef (SE)	Coef (SE)	Coef (SE)
	RCV	RCV	RCV
Within-herd BVDV transmission rate by a PI animal (β_P)	14.92 (0.06)	0.63 (0.01)	-122,842.6 (2,214.0)
	22.2%	20.5%	11.4%
Relative proportion of β_B to β_P	0.17 (0.01)	0.01 (0.00)	-2,494.3 (280.0)
	0.2%	0.5%	0.3%
Relative proportion of β_F to β_P	0.05 (0.01)	0.00 (0.00)	-478.9 (280.0)
	0.0%	0.0%	0.0%
Prevalence of PI animals in the neighbouring farms (K_F)	1.03 (0.08)	0.06 (0.01)	-9,446.1 (2,800.5)
	0.0%	0.1%	0.0%
Mortality of PI animals (μ)	-675.19 (1.51)	-29.13 (0.13)	-4,280,474.1 (56,336.3)
	70.3%	68.5%	21.4%
Number of introduced Trojan cows as a relative proportion to T_M (ρ)	8.47 (0.06)	0.45 (0.01)	-297,288.0 (2,214.0)
	7.2%	10.4%	66.8%

Key: *TI*, transiently infected animal; *PI*, persistently infected animal; *GFR*, gross farm revenue; Coef, coefficient; SE, standard error of the coefficient; RCV, relative contribution of output variation.

Appendix 4

Section 1. Modelling of animals on beef fattening or hobby farms

Due to the lack of data about the management characteristics of individual beef fattening or hobby farms, the demographic and BVDV transmission status of animals on each of these farms could not be explicitly modelled. To address this, we generated a virtual location that all dairy-beef animals (i.e. fattening beef cattle with dairy origin) were moved at 1 week old from every simulated dairy farm, and assumed that the animals on the virtual location represented the animals on all beef fattening and hobby farms across New Zealand. Same as the beef fattening stock on beef breeding farms, we assumed that female and male dairy-beef animals were slaughtered when they reached 118 or 98 weeks old, respectively. BVDV transmission among dairy-beef animals in the virtual location was also assumed with the probability of BVDV infection for a susceptible dairy-beef animal on a given week (P_{inf}^{db}) as;

$$P_{inf}^{db} = 1 - e^{-\lambda^{db}}; \lambda^{db} = \beta_P \rho_B$$

where ρ_B was the prevalence of PI animals in the virtual location.

Since the animals on beef fattening or hobby farms were implicitly modelled, the impact of those farms on national BVDV transmission was indirectly incorporated into the national model by assuming that (1) the prevalence of PI animals on all beef fattening and hobby farms across New Zealand was the same to ρ_B , and (2) adjacent beef fattening and hobby farms contributed to local spread on dairy or beef breeding farms. A beef fattening or hobby farm was assumed be adjacent to a given dairy or beef breeding farm if the Euclidean distance between two farms were ≤ 3.0 km (see Section 3 in Appendix 4). With these assumptions, we treated ρ_B as the background prevalence of PI animals in New Zealand.

Section 2. Modelling of animals on beef fattening or hobby farms

2.1. Inferring herd size distributions

2.1.1. Dairy farms

We assumed that the distribution of number of mixed-age cows (MA) in New Zealand dairy farms followed $e^{N(\mu_D, \sigma^2)}$, where $N(\cdot, \cdot)$ was a normal distribution, and μ_D and σ were unknown mean and standard deviation, respectively. We used an approximate Bayesian computation-sequential Monte Carlo (ABC-SMC) method to infer the unknown μ_D and σ . ABC-SMC is a Bayesian method to infer the parameter values through a series of estimation sequences (or SMC sequence), each of which has a decreased distance threshold than its previous sequence (Toni et al., 2009). A SMC sequence starts with selecting a parameter value (e.g. μ_D and σ) based on the perturbation of estimated values in the previous sequence (or the random sampling of a prior distribution in case of the first SMC sequence), and applying the selected value to a model (i.e. $e^{N(\mu_D, \sigma^2)}$) to generate data (i.e. simulated data). The difference between observed and simulated data is measured by comparing the summary statistics of each data (i.e. distance), and only the value showing the distance less than a threshold of the sequence is retained. Current SMC sequence ends when a number of values (i.e. particle) are collected, and the whole process repeats until the particles of final SMC sequence are collected. The distribution drawn by the particles of final sequence is a marginal posterior distribution, which approximates the posterior distribution of the parameter given the observed data. More detailed information about ABC-SMC method can be found in elsewhere (Toni et al., 2009).

In this study, a random MA herd size was generated based on the selected μ_D and σ for each dairy farm, and the farm was allocated to one of 24 intervals of MC herd size ((0,50], (50,100], (100,150], ..., (950,1000], (1000,1100], (1100,1200], (1200,1500], [1500,2000)) according to the generated herd size. The distance was measured once all MC herd sizes of 11,269 farms were allocated to the relevant interval, and the selected μ_D and σ were either

kept or discarded depending on the distance threshold value of each SMC sequence. The distance between summary statistics of observed and simulated data was measured as;

$$\sqrt{\sum_{i=1}^n (Obs_i - Sim_i)^2}$$

where n was 24 which was the total number of MA herd size intervals, Obs_i and Sim_i were the proportion of farms allocated to interval i over total number of dairy farms in observed and simulated data, respectively.

The number of dairy farms in each interval of MA herd size was provided in Anonymous, (2018a). We collected 2,000 particles for each SMC sequence, and μ_D and σ were searched for total number of 15 SMC sequences. Prior distribution of μ_D and σ were arbitrarily chosen as $U(1, 8)$ and $U(0, 4)$, respectively, where U indicated an uniform distribution. For the perturbation kernel, a component-wise Gaussian kernel with the variance as 0.68 times of the variance of particles in the previous SMC sequence was used (Ellen Brooks-Pollock et al., 2014). The acceptable distance thresholds were set as 50th percentile of the distances of accepted particles in the previous sequence.

The posterior distribution of μ_D and σ is provided in Figure S8 and S9, respectively. The most probable (mode) μ_D and σ values were 5.83 and 0.61, respectively. The simulated distribution of MA herd sizes in New Zealand dairy farms based on the most probable μ_D and σ values is illustrated in Figure S10 with the observed distribution for comparison.

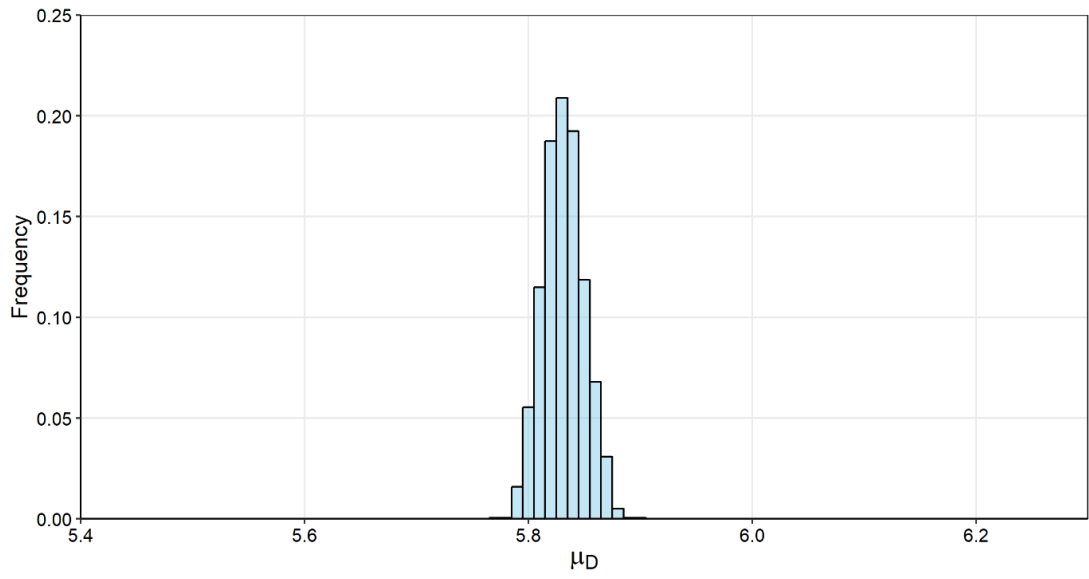


Figure S8. Posterior distribution of μ_D .

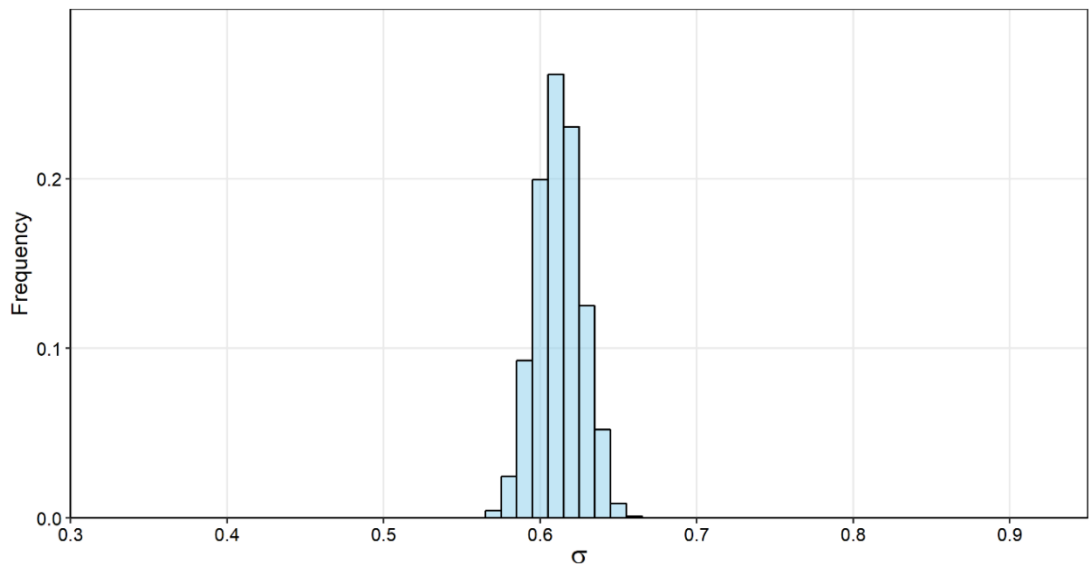


Figure S9. Posterior distribution of σ .

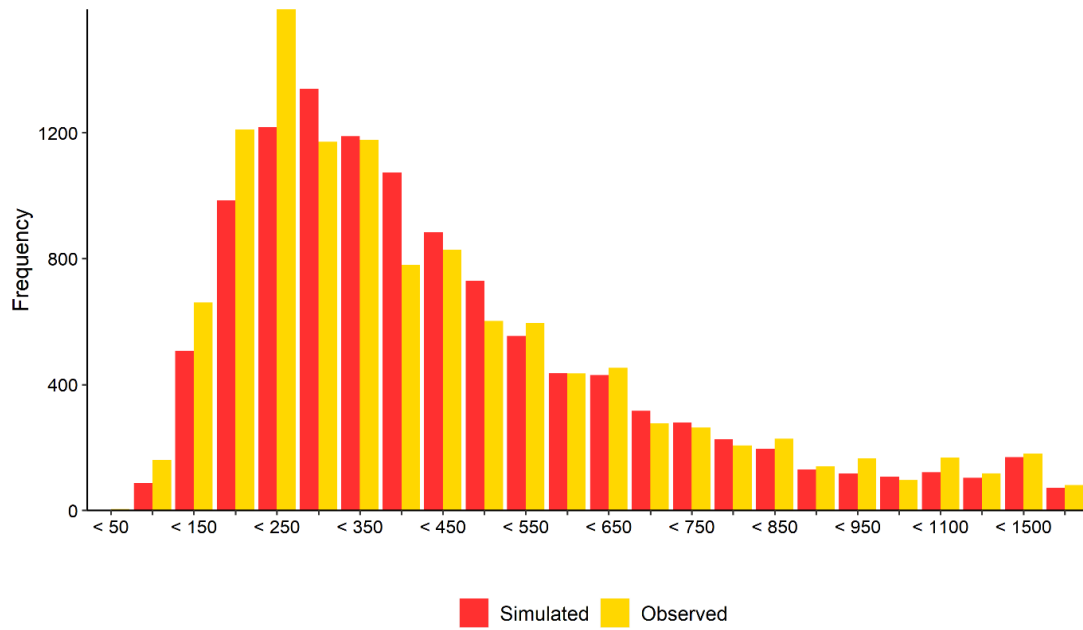


Figure S10. Observed and simulated distribution of mixed-age herd size of 11,269 dairy farms in New Zealand.

2.1.2. Beef breeding farms

Unlike the dairy sector where a brief illustration of national MA herd size exists, there was no equivalent information in the beef sector in New Zealand. The only information available was the total number of MAs in New Zealand beef farms (approximately 850,000) (Anonymous, 2018c). Therefore, we assumed that the distribution of MA herd size in beef farms followed $e^{N(\mu_b, \sigma^2)}$, where σ was 0.61 which was estimated in the dairy sector. We used an ABC-SMC method to infer the unknown μ_B while restricting the MA herd size as bigger than or equal to 10. A random MA herd size was generated based on the selected μ_B and σ for all 4,654 beef farms, and we calculated the total number of MA cows which was the summary statistics. Details of the ABC-SMC method was the same as in the dairy farms, except the distance was measured as the absolute difference between the observed and simulated summary statistics.

The posterior distribution of μ_B is provided in Figure S11, and the most probable (mode) μ_B was 5.02. The simulated distribution of MA herd sizes in New Zealand beef farms is

provided in Figure S12. Unfortunately, we were not aware of any data or reference to validate the simulated distribution of MA herd size of beef breeding farms in New Zealand. However, the median of the simulated distribution was 150, which was in great accordance with the reported median herd size of beef MA from a study in New Zealand (McFadden et al., 2005).

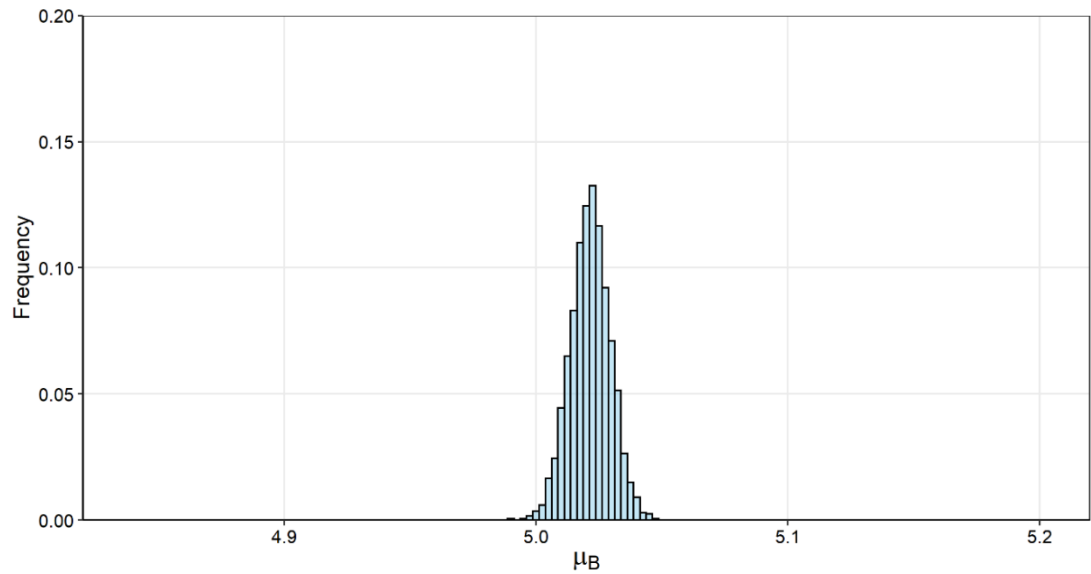


Figure S11. Posterior distribution of μ_B .

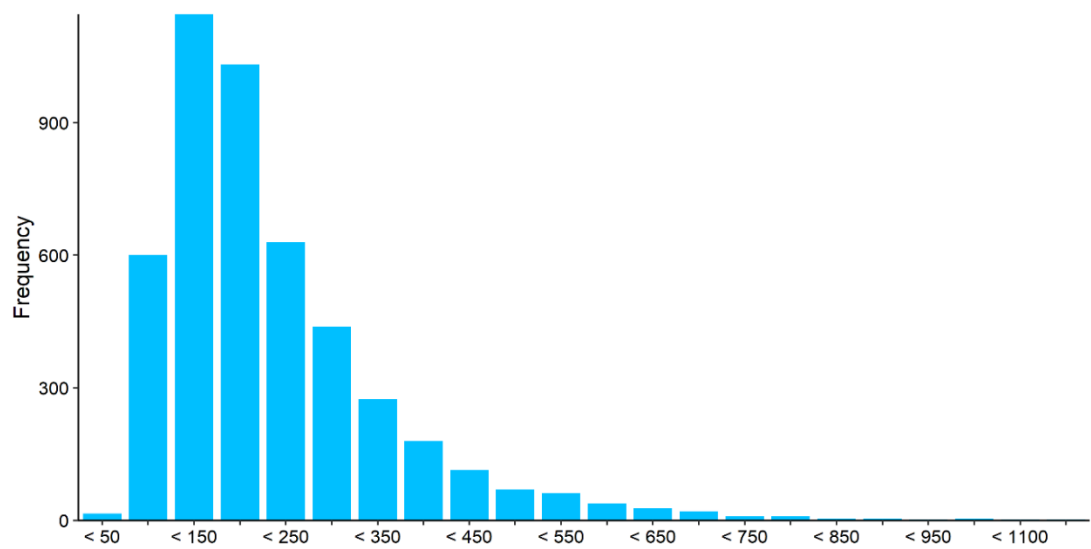


Figure S12. Simulated distribution of mixed-age herd size of 4,654 beef breeding farms in New Zealand.

2.2. Allocation of a point location for farms without a geocoordinate

Among the active cattle farms, 1,452 (12.9%) of 11,269 dairy farms, 572 (12.3%) of 4,654 beef breeding farms, 609 (13.3%) of 4,573 beef fattening farms, and 2,785 (11.0%) of 25,320 hobby farms did not have their geo-location information. We allocated the coordinates of those farms by (1) dividing whole New Zealand into 1,000 by 1,000 grids, of which each cell contained the proportion of cattle farms with respective production type (with coordinates) in the cell over entire cattle farms with respective production type (with coordinates), (2) allocating a farm without coordinates to a random cell while applying the proportion of each cell as a weight for the random selection, and (3) randomly assigning coordinates inside the allocated cell. Once the allocation process was done, we confirmed that single farm had unique geo-coordinates.

2.3. Description of demographics within cattle farms

2.3.1. Dairy farms

Cattle population for a dairy farm was divided into 4 management groups; calves (C: until 1 week old), young replacement heifer (YH: from 1 week old to off-grazing), breeding replacement heifers (BH: heifers at the off-grazing location), and mixed-aged cows (MA). All new born calves were separated from their dams as soon as they were born and joined the C group. The first x number of female calves that reached 1 week old were retained as YHs where x was approximately 20.0% of the assigned MA herd size. Among the remaining female calves and all male calves, 55.0% of them were sent to abattoirs (i.e. bobby calves) and the rest were sent to a virtual location for dairy-beef cattle (see Section 1 in Appendix 4). The proportion of bobby calves was calibrated to match the annual total number of bobby calves culled to be approximately 1.85 million. To adjust for the natural mortality, 90.5% animals were assumed to survive through weaning (at 10 weeks old) (E. Cuttance et al., 2017). YHs were sent for off-site grazing on a random week between week

18 and week 22, and returned on the same week of the following year. Among returned BHs, only the pregnant heifers were mixed to the MA group while the rest were culled.

Puberty in heifers was assumed to start at 54 weeks old (McNaughton et al., 2002) with oestrus cycles occurring every 3 weeks thereafter until the animal became pregnant. When a pregnant animal calved or aborted, the next oestrus was assumed to occur after two to seven weeks (Olds & Seath, 1953) and was assumed to be silent. For simplicity, the planned start of mating (PSM) for both the MA and BH herds started on the same week (randomly chosen between week 39 and 45) and lasted for 12 weeks. During the first 6 weeks, animals in oestrus were artificially inseminated. The probability of oestrus detection was set at 78.6% and the probability of conception to artificial insemination was set at 72.0%. For the last 6 weeks, animals in oestrus were naturally mated with breeding bulls with the conception probability at 50.0%. These probabilities were calibrated based on 66.1% of cows artificially bred for the first 6 weeks being pregnant and 15.0% of cows remaining non-pregnant after the 12 weeks of mating period (Anonymous, 2018a).

The gestation period was set at 40 weeks. A natural abortion rate of 0.0007 per week was applied over the whole gestation period to achieve an overall abortion probability of 3.5% (Weston et al., 2012). Pregnancy scanning was conducted after 18 weeks from the PSM for MAs only. Non-pregnant cows were culled immediately unless they were lactating, in which case the cows were scheduled to be culled at the start of drying-off period. For simplicity, we assumed that the start of drying-period coincided with the week of off-site grazing. We assumed that only the pregnant BHs joined the MA herd, and we kept the MA herd size constant by (1) culling x number of pregnant MAs older than 6 years where x was the surplus of MAs compared to the assigned MA herd size, or (2) keeping x number of empty MAs where x was the shortage of MAs compared to the assigned MA herd size. The herd of breeding bulls was not considered in the simulation model since it is common practice for breeding bulls to be present on farm only for the duration of the mating period

and for most breeding bulls to be tested for, and vaccinated against, BVDV (Han, Holter, et al., 2018).

2.3.2. Beef breeding farm

For the simulation model, we divided the beef population into 7 management groups; Fattening heifers (FH) and steers (FS) for finishing, young replacement heifers (YH: heifers in this herd are bred for the first time), breeding replacement heifers (BH: heifers in this herd calve for the first time and are bred for the second time) and their calves (CH), and mixed-age cows (MA) and their calves (CM).

The start of puberty, oestrus cycle, gestation period, and natural abortion rate were assumed to be the same as dairy heifers/cows (Hickson et al., 2012). The PSM was randomly chosen between week 38 and 52, and cows/heifers were grazed with breeding bulls for 9 weeks of mating period. For the animals in oestrus, the probability of conception by naturally mating with breeding bulls was 72.0% (Geenty & Morris, 2017). Off-site grazing was not considered in the beef model since it is not a common practice in the beef industry.

As beef cattle in New Zealand are infrequently handled compared with dairy cattle, we assumed that most of the transitions between demographic management groups occurred on the week of weaning, which was 19 weeks after the PSM. We also assumed that the natural mortality rate of new born calves differed with 86.7 and 95.0% of calves born from BH and MA, respectively, surviving until weaning (Hickson et al., 2012; C. A. Morris et al., 1986). When the calves were weaned, up to x number of female calves were selected among those who were born during the first 3 weeks of the calving period and transferred to the YH group, where x was approximately 20.0% of the assigned MA herd size. Other calves were transferred to either FH or FS depending on their sex, and culled when they reached 118 or 98 weeks of age, respectively (Geenty & Morris, 2017). Transition of YHs

into BHs also occurred on the week of weaning, and it was assumed that all BHs joined the MA herd at 3 weeks before the week of weaning. At one week before the week of weaning, pregnancy scanning of MAs was performed and any empty MAs were tagged. We kept the herd size constant by either (1) selling some of the pregnant (i.e. non-tagged) MAs older than 6 years if the number of MAs (= non-tagged MAs + joined BHs) was larger than the assigned MA herd size, or (2) keeping empty (i.e. tagged) MAs if the number of MAs (= non-tagged MAs + joined BHs) was lower than the assigned MA herd size. As in the dairy model, the demographic group of breeding bulls was not included in the beef model.

2.4. Probability of having a cattle trade between two farms

We reconstructed the cattle trade movements between dairy farms and between beef breeding farms using a network re-wiring algorithm based on the observed trade data during 2014-2016. The network re-wiring algorithm is a method to connect a pair of farms based on the probability of whether a trade was likely to occur between two farms.

To predict the probability of a trade between two dairy (or beef breeding) farms, we first developed a multivariable logistic regression model. Using the extracted trade data, we generated: (1) a list of farms sold and purchased cattle, and (2) the number of animals sold and purchased per farm, for every week. Then, we created a contingency dataset that contained all possible dairy (or beef breeding) cattle trades that could have occurred. For example, if actual trades occurred between farms A and B, C and D, and E and B (sold and purchased) in a given week, the contingency dataset was created with the trades between A and B, A and D, C and B, C and D, E and B, and E and D, while only the ones actually occurred were tagged as positive for the outcome of the logistic regression. For each trade in the contingency dataset, the Euclidean distance, (absolute) difference in the number of cattle traded, and whether two farms were located in the same bovine tuberculosis control area were investigated, and used as the explanatory variables of univariable logistic regression. For the Euclidean distance and difference in the number of cattle traded,

different transformations — $\ln(x)$, $\frac{1}{e^x}$, and x exponentiated with -3, -2, -1, -1/2, -1/3, 1/3, 1/2, 1, 2, and 3 — were tested to adjust the non-linear relationship, and the one showing the lowest AIC in the univariable logistic regression model was incorporated in the multivariable logistic regression. We only considered the full model which contained all the (transformed) explanatory variables in the multivariable logistic regression.

For the trade of both dairy and beef cattle, the transformation of $\ln(x)$ and $\frac{1}{e^x}$ showed the lowest AIC value for the variable of Euclidean distance and the difference in the number of cattle traded, respectively. Table S6 provides the results of multivariable logistic regression models. Overall, 27.4% of variance was explained by the explanatory variables in the trade of dairy cattle, whereas it was 36.7% in the trade of beef cattle.

Table S6. Results of multivariable logistic regression models for predicting the probability of cattle trades between farms.

Variable	Dairy farms			Beef breeding farms		
	Coef	SE	P-value	Coef	SE	P-value
Constant	-1.13	0.02	<0.001	1.75	0.06	<0.001
Euclidean distance (km) *	-1.08	0.00	<0.001	-1.20	0.01	<0.001
Difference in the number of cattle traded **	2.42	0.01	<0.001	2.80	0.04	<0.001
BT movement restriction						
From NR to NR	Ref.		<0.001	Ref.		<0.001
From NR to R	0.57	0.04	<0.001	0.36	0.07	<0.001
From R to NR	0.57	0.04	<0.001	0.41	0.07	<0.001
From R to R	1.53	0.07	<0.001	0.48	0.09	<0.001

Key: Coef, coefficient; SE, standard error of the coefficient; BT, bovine tuberculosis; NR, not-restricted; R, restricted; Ref, reference.

* The variable was transformed with $\ln(x)$.

** The variable was transformed with $\frac{1}{e^x}$.

Section 3. Calibration of parameters related to between-farm BVDV transmission

In this study, we assumed that between-herd transmission rate (β_H) and between-farm transmission rate (β_F) of a PI animal were proportionally (m_1 and m_2 , respectively) lower than within-herd transmission rate of a PI animal (β_P). We calibrated the values of m_1 and m_2 as well as the distance threshold to determine adjacency between farms (d) by comparing the simulated prevalence of BVDV positive farms to the reported prevalence.

To estimate the parameter values while achieving computational efficiency, we developed a sub-model which included only the cattle farms (1,212 of dairy farms and 873 of beef farms) in the Canterbury. We selected Canterbury region since it contained approximately 10% of all commercial cattle farms in this study. Trade movements between farms in Canterbury region were extracted to estimate the total number of cattle traded per week and total number of farms sold and purchased cattle per week inside the region, and we ignored any trade from (or to) farms in the region to (or from) farms in other regions.

In the sub-model, we ran the first 5 years as a burn-in period to stabilise the population demographics, introduced BVDV to randomly chosen 104 farms (5% of farms in the Canterbury), and simulated the model for the next 40 years to measure the proportion of dairy farms carrying at least one either TI or PI animals in mixed-age cow herd ($Prev_+^D$). The $Prev_+^D$ was a proxy of the proportion of positive farms to bulk tank milk (BTM) BVDV antigen ELISA or RT-PCR test. We assumed that the median of $Prev_+^D$ during the last 15 years should approximate at 7.5% (Gates, Han, et al., 2019), and identified the m_1 , m_2 , and d values that satisfying the assumption by (1) applying different combination of values (m_1 , and m_2 : 0.10, 0.20, and 0.40 for each; d : 2.0 km, 3.0 km, and 4.0 km) into the sub-model (resulting in 27 scenarios), and (2) repeating the (1) process for 50 times to generate the distribution of $Prev_+^D$ with respect to different combinations of m_1 , m_2 , and d values. We then visually examined the generated distributions, and chose the parameter values that the median of the distribution was located approximately 7.5%. If the median of $Prev_+^D$ for the multiple combinations of m_1 , m_2 , and d values approximated to the 7.5%,

the combination that the median of $Prev_+^D$ being the most close to 7.5% while its $m_1 \geq m_2$ was chosen given the higher chance of between-herd transmissions than between-farm transmissions due to the extensive farming condition in New Zealand.

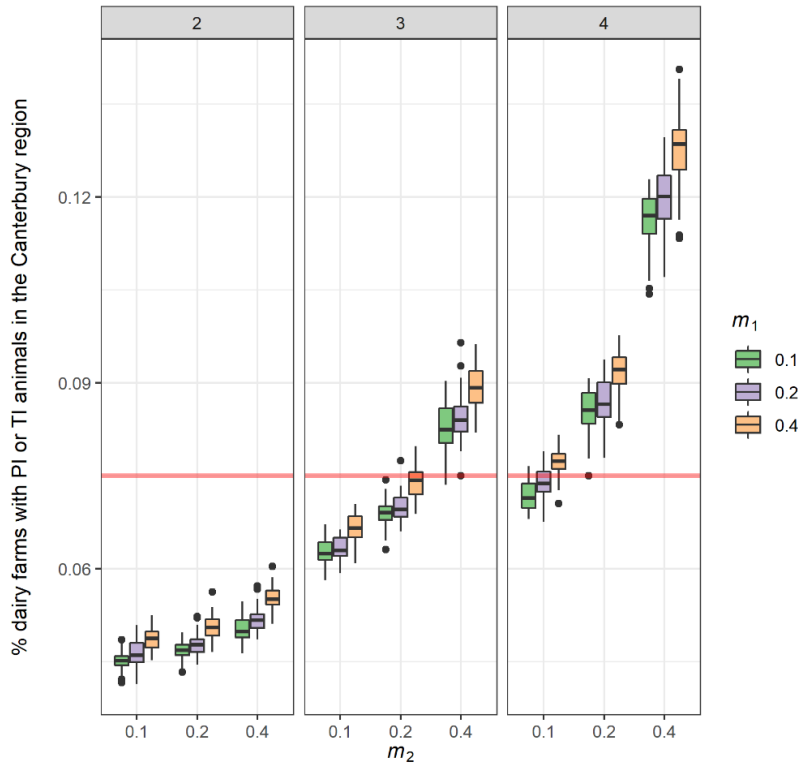


Figure S13. Prevalence of dairy farms carrying at least one either TI or PI animals in mixed-age cow herd in the Canterbury region. The prevalence varied with respect to different between-herd transmission rate (m_1), between-farm transmission rate (m_2), and the distance threshold for adjacency of cattle farms (d). Values in grey-shaded area at the top of each column indicate the value of d , and the red line indicates the prevalence of positive dairy farms to bulk tank milk (BTM) BVDV antigen ELISA or RT-PCR test in New Zealand.

The distributions of $Prev_+^D$ in the Canterbury region with respect to different parameter values are illustrated in Figure S13. The median of $Prev_+^D$ for the two combinations of m_1 , m_2 , and d values (0.40, 0.20, and 3.0 km, respectively; 0.20, 0.10, and 4.0 km, respectively) approximated to the 7.5%. Among these, we prefer to use 0.40, 0.20, and 3.0 km for m_1 , m_2 , and d , respectively. This was because the maximum number of adjacent farms was over 120 with d value of 4.0 km (data not shown), and we believed that using the latter

combination (0.20, 0.10, and 4.0 km for m_1 , m_2 , and d , respectively) would overestimate local spread of BVDV transmission. Therefore, we concluded the value of β_B and β_F as 0.40 and 0.20 times of β_p , respectively, and the distance threshold for adjacency of cattle farms as 3.0 km well represented the parameters for BVDV transmission in the New Zealand cattle industries.

Section 4. Overall modelling process

At the beginning of each week, farms were randomly selected as either source or destination farms for animal trade movements, and paired based on a network re-wiring algorithm. For each farm of a given week, the probability of BVDV infection for individual animals was calculated for each herd initially. Then, demographic and BVDV infection status of each individual animal in each herd were updated. Once the demographic and BVDV infection status were updated, the animal was probabilistically determined to be whether traded or not if the farm was on the list of source farms. After the simulation of all animals for all farms was completed for the given week, the model moved on to the next week until the end of simulation period. The overall modelling process is illustrated in Figure S14.

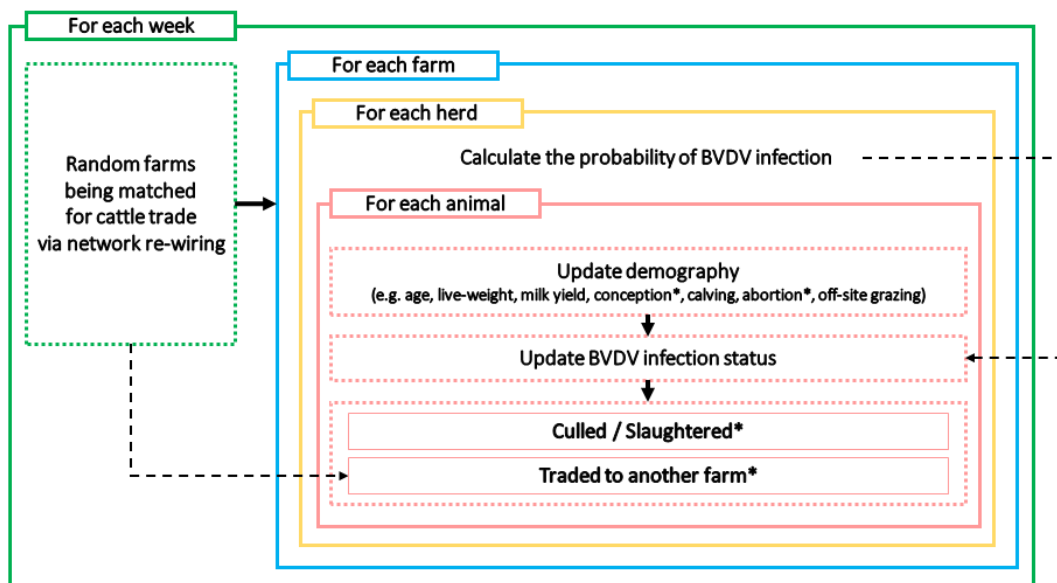


Figure S14. Flowchart of the overall modelling process. * indicates the demographic events that occurred probabilistically.

Section 5. Justification for the number of simulations

Given the high computational cost of running the national BVDV simulation model, we restricted the number of iterations for each control programme to 100 times. To validate whether the number of iterations for model simulation adequately captured the variation caused by the stochasticity, we examined the impact of number of iterations on the variation of simulated national BVDV prevalence (i.e. proportion of farms carrying at least one PI animal). First, we ran the national model without any control programmes for 100 times to store the national BVDV prevalence over the final year of simulation periods (year 45). We then (1) selected the national BVDV prevalence from randomly selected 5 iterations and measured the standard deviation, (2) repeated the process of (1) for 100 times to estimate the interquartile range (IQR) of the standard deviation of national BVDV prevalence from 5 iterations, and (3) repeated the whole process of (1) and (2) with the increment of 5 random iteration sizes until it reached 50. The estimated IQR of the standard deviation of national BVDV prevalence with varying iteration sizes was visually inspected. For convenience, we only examined the variation of national BVDV prevalence for week 1, week 18, and week 35 of the final years.

Figure S15 illustrates the IQR of the standard deviation of national BVDV prevalence with respect to the number of iterations for simulation over different time points. It shows that the IQR noticeably reduced until 20 iterations, and then plateaued at 50 iterations. This indicates that the variation of national BVDV prevalence that caused by the lack of iterations substantially reduced by 50 iterations of simulation, and the variability caused by stochasticity was adequately captured with only 100 iterations.

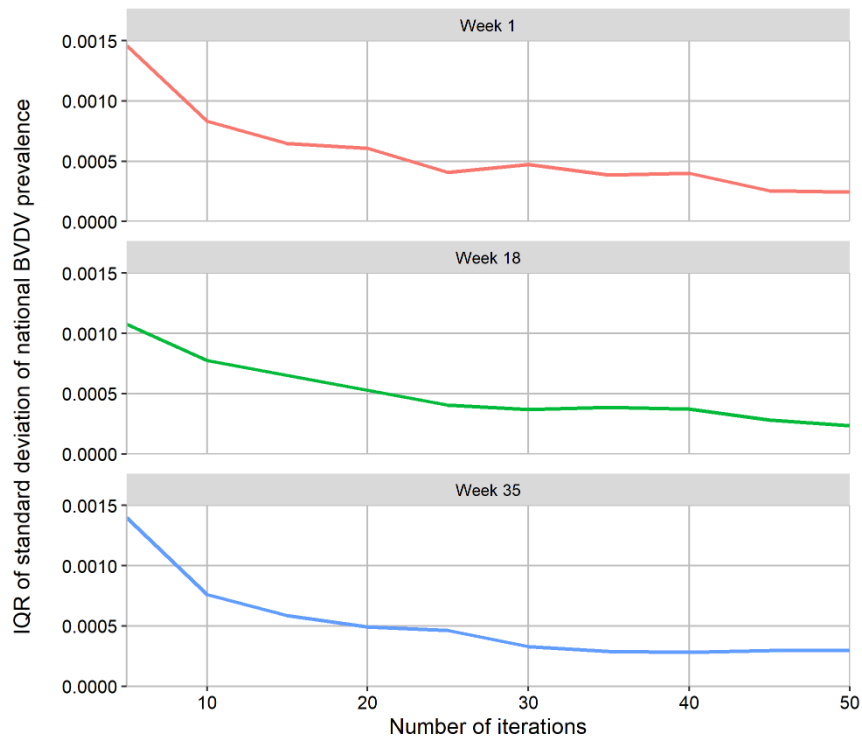


Figure S15. Interquartile ranges (IQR) of the standard deviation of national BVDV prevalence with respect to the different number of iterations for simulating a national BVDV simulation model. IQRs were measured in three different time points of the final year of simulation period.

Section 6. Correlation between the herd size and BVDV endemic status

Since double fencing could not completely eliminated BVDV in the simulated beef breeding farms in the national model, we analysed the association between the herd size and BVDV endemic status of the beef farms when double fencing was implemented. We defined farms were BVDV endemic if any PI animals presented in the farms at the end of simulation period. For each beef breeding farm, we then inspected the proportion of iterations that the farm was BVDV endemic over 100 total iterations. Correlation between the herd size of mixed-age cow (MA) and the proportion of endemic BVDV was visually examined.

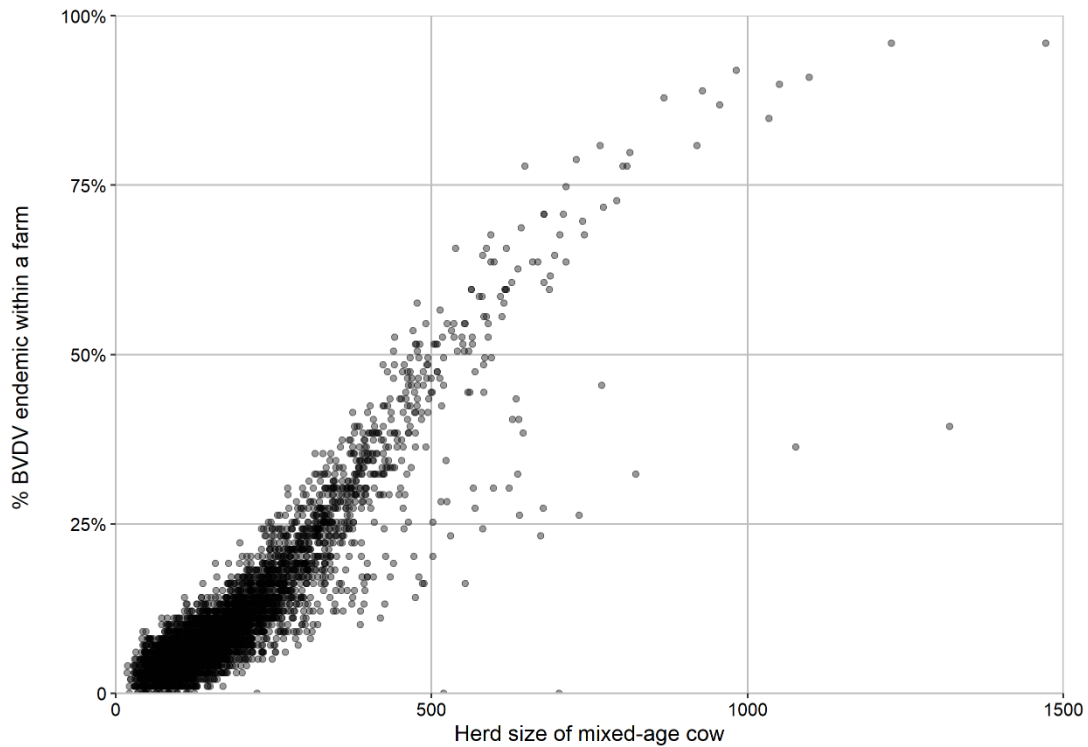


Figure S16. Correlation between the herd size and proportion of endemic BVDV in 4,654 simulated beef breeding farms.

Figure S16 illustrates the association between the herd size and proportion of BVDV endemic status for 4,654 beef breeding farms. BVDV persisted until the end of simulation period in approximately 50% of iterations if farms had more than 500 MA herd size. It shows that the endemic status of simulated beef breeding farms was highly correlated with their herd size.



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Name of candidate:	Jun-Hee Han
Name/title of Primary Supervisor:	Carolyn Gates
Name of Research Output and full reference:	
Elimination of bovine viral diarrhoea virus in New Zealand: a review of research progress and future directions	
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JH Han reviewed the literature and drafted the manuscript.	
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JH Han established the model, analysed results, and drafted the manuscript.	
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