

Article

A New Look on Financial Markets Co-Movement through Cooperative Dynamics in Many-Body Physics

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Abstract: One of the main contributions of the Capital Assets Pricing Model (CAPM) to portfolio theory was to explain the correlation between assets through its relationship with the market index. According to this approach, the market index is expected to explain the co-movement between two different stocks to a great extent. In this paper, we try to verify this hypothesis using a sample of 3.000 stocks of the USA market (attending to liquidity, capitalization, and free float criteria) by using some functions inspired by cooperative dynamics in physical particle systems. We will show that all of the co-movement among the stocks is completely explained by the market, even without considering the market beta of the stocks.

Keywords: econophysics; collective motion; finance; stock market; capital assets pricing model

1. Introduction

The Capital Asset Pricing Model (CAPM) by Sharpe [1] Lintner [2] and Black [3,4] is one of the most relevant models within financial economics. The contribution of the CAPM to finance was associating the return on a financial asset with its systematic risk, which is measured by the market parameter Beta.

The CAPM approach to the study of the return on financial assets is as follows:

$$r_i - R_f = \alpha_i + \beta_i(R_m - R_f) + \epsilon_{it} \quad (1)$$

where,

r_i is the return on the asset i ,

α_i is the part of the return on the asset i that is due to aspects of the asset itself, such as profit, business, etc.

R_f is the return of the risk-free asset,

R_m is the market return (usually proxied through its main equity index or an all-shares value-weighted portfolio),

$(R_m - R_f)$ is known as the market risk premium, and

ϵ_{it} is a zero-mean residual.

In a sense, the CAPM is used to explain the differences in risk premiums for different assets in the market, which are caused by differences in the risk of asset returns. According to the CAPM, β_i is the factor that measures the sensitivity of the return on asset i to market movements. Accordingly, if the risk premium of a particular asset needs to be predicted, it is only necessary to select a risk-free rate and the beta of this asset. This model has been widely studied not only in the financial literature, but also by practitioners. However, there is a permanent debate about the usefulness of the market model, since, over time, empirical work has emerged with contrary results about whether market beta is indeed a useful factor in the study of a portfolio's performance. Blume [5], Black et al. [6], and Fama and Macbeth [7] performed the first tests over the CAPM and stated that indeed it provides a good prediction model. Conflicting results were obtained by King [8], who found the existence of variance due to industry effects and that the market effect accounts for only about 50 per-cent of the variance. Meyer [9] also found numerous unexplained components that represent some persistent significant source of interdependence among stock prices. Banz [10] found that the size effect of firms explained better than market *beta* the variation in average returns at a portfolio. Roll [11] found that market *beta* explains monthly stock price changes only in an average of 30%. However, testing other factors, such as industry or size, this author concluded that explanatory power by systematic economic factors is not different across industries. Bhandari [12] also found a positive relationship between leverage and average return. This author concluded that *beta* leaves out important information in order to explain more accurately the reality. Based on previous results of Stattman [13] and Rosemberg et al. [14], Fama and French [15] proposed an alternative model including three factors: the market *beta*, the size effect and the book to market ratio. According to their results, the market *beta* barely had the capacity to explain the variations in the average yields of the selected stocks, while size and value factors showed interesting results. The new factorial model proposed by Fama and French made clear the insufficiency of the market *beta* to cover all of the information needed to explain the variations in returns. Over time, advocates of the market *beta* have been studying the reasons why the CAPM no longer performs well in empirical studies. A good example is found in Isakov [16]. In this paper, the authors stated that *beta* has no chance of being a useful variable in recent empirical studies for two reasons. The first reason is that the model is expressed in terms of expected yields, but the tests can only be performed on yields that have been realized. The second reason is that excess realized market returns do not behave as expected, i.e., they are too volatile and often negative. There are similar works with different methodologies proposed, where it is shown that the market beta is still valid under some circumstances (see, for example, Pettengill, Sundaram, and Mathur [17], Chan and Lakonishok [18], Grundy and Malkiel [19], or Lopez et al. [20])

For our purposes, one of the main goals of the CAPM is to explain the correlation of different assets through its relationship with the market index given by Equation (1). This represented an important advance in portfolio selection, because the market index is expected to explain all of the co-movement between two different assets, leaving only the residual ϵ_{it} to differ the trend of asset prices.

Assets co-movement has captured the attention of researchers due to its importance for asset allocation, portfolio diversification, or risk management. Causes of co-movement have been studied from different points of view. Roll [11] found that the level of stocks co-movement depends on the relative amounts of firm-level and market-level information capitalized into stock prices. Domowitz et al. [21] proved that liquidity co-movement is determined by order flow and return co-movement is caused by order type (market and limit orders). Byrne et al. [22] found that co-movement is responsible for two-thirds of variability in global bond yields. These authors concluded that global inflation explains most of the global yield co-movement.

Morck et al. [23] found that stock returns are more synchronous in emerging markets. The authors also show evidence that co-movement is not a consequence of structural characteristics of economies, such as fundamentals volatility or country and market size. Jach [24] quantified time-varying, bivariate, and multivariate co-movement between international stock market return. This work concluded that development and region are not always decisive factors. Parsley and Popper [25] observed that return

co-movement does not depend on the richness of the country, but it is more affected by variables that reflect different institutional aspects, including international macroeconomic policy stability.

Other authors have focused their attention on the co-movement between different assets or countries. Bonfiglioli and Favero [26] did not find evidence of long-term interdependence between US and German stock markets. Rua and Nunes [27] analyzed the co-movement across major developed countries among 40 years. These authors found that the co-movement is strong across countries in lower frequencies and it is different in each one. They also remarked that the degree of co-movement varies over time. Akoum et al. [28] examined the co-movements of stock markets in the GCC region and crude oil prices. These authors showed evidence of a strong dependence after 2007 in the long term. However, in the short term, there is no clear evidence of dependence. Reboredo [29] found no extreme co-movement between oil prices and exchange rates in the periods before and after the financial crisis. Magdaleno and Pinho [30] reported a strong and significant relationship between index prices. The authors show that innovations in the US and UK stock markets are not rapidly transmitted to other markets. They also found that, economically, as well as geographically, economies show higher levels of co-movements, except Japan. Loh [31] studied the co-movement of 13 Asia-Pacific stock markets and the European and US ones. These authors found a significant co-movement between most of them in the long run. During the financial crisis, this author reports evidence of important variation in co-movement in time.

Factors relevant in co-movement have also been analyzed by some authors. Baca et al. [32] showed evidence of the importance of global industry factors in explaining international return variation. In a similar line, Cavaglia et al. [33] showed that country effect is more relevant than industry factors in the late-1990s and Griffin and Karolyi [34] found that global industry factors only explain four percent of the variation in local stock markets. However, L'Her et al. [35] reported evidence in an opposite sense, finding that global industry effects surpassed country effects in importance in 1999–2000. Brooks and Del Negro [36] analyzed the link between international stock market co-movement and firm level variables that measure international diversification, finding a significant relationship between stock returns *betas* and those variables. Antonakakis and Chatziantoniou [37] showed that there is a clear negative correlation between the policy uncertainty and stock market returns, except during the financial crisis.

Different approaches have been used to measure co-movement, such as cross-correlation analysis (Akoum et al. [28]), spatial techniques (Fernández-Aviles et al. [38]), regression coefficient (Brooks and De Negro [36]), quantiles (Cappiello et al. [39]), tail-dependence coefficient (Garcia and Tsafack [40]), copula approach (Reboredo [29]), time series analysis (Antonakakis et al. [37]), shortfall-multidimensional scaling approach (Fernández-Aviles et al. [41]), or a recent approach based in Hurst Exponent (Ramos Requena et al. [42]). In this paper, we propose a new approach to study the co-movement of the whole market based on some functions borrowed from or based on physical particle systems, following our previous works on this topic (Clara Rahola et al. [43], Sánchez Granero et al. [44], Puertas et al. [45]).

From a physical point of view, a portfolio can be seen as a many-body system, with the index representing a particular point of the system characterizing the whole system, such as the center of mass. The main contribution to the asset motion is described by Brownian motion, and the interactions among the assets are unknown, if they exist, and they induce collective motion. However, as given by classical mechanics of many-body systems, internal forces only affect the relative motion of a body with respect to the center of mass. Thus, a proper account of cooperative motion, within a physical approach can only be made if the center of mass (or index) is subtracted.

We study co-movement in the whole market in this paper by considering some functions inspired by co-movement in physical particle systems. It is showed that the market (represented by an equal weighted portfolio of all stocks) explains all of the co-movement in the whole market. Alternative ways to represent the market are considered, for example, portfolios weighted by capitalization or known

indexes, in order to check whether these representations are also able to explain all the co-movement in the market. Some other alternatives that take into account the market beta are also considered.

2. Methodology

Cooperative dynamics have been observed in many physical particle systems. Furthermore, due to the short-time Brownian motion of assets, here we focus on particles with Brownian dynamics, namely colloidal systems. In colloids, as well as in atomic or molecular fluids, upon lowering the temperature the dynamics of the system slows down until the glass transition is reached, and the dynamics freezes, ideally. Accompanying this slowing down, the dynamics becomes more cooperative with more mobile particles clustering together in string-like structures (Donati et al. [46], Cates and Evans [47], Weeks et al. [48]). To study this phenomenology, several so-called four points correlation functions have been proposed, relating the dynamics of two particles at two different times (Glotzer et al. [49], Berthier et al. [50]). The most direct one is the evolution of the distance between a given pair of particles (Muranaka et al. [51], Zahn et al. [52]). Inspired by these functions, here we propose three simple observables that can be used to monitor the cooperative dynamics in financial systems:

- C_1 . Consider the functions

$$C_{1t}(\tau) = \frac{\sum_{i,j} (\delta x_i(t, \tau) - \delta x_j(t, \tau))^2}{\sum_{i,j} (\delta x_i(t, \tau)^2 + \delta x_j(t, \tau)^2)}$$

$$C_1(\tau) = \langle C_{1t}(\tau) \rangle$$

where $x_i(t)$ is the log-price of asset i at time t , $\delta x_i(t, \tau) = x_i(t + \tau) - x_i(t)$ and the sum is for each pair of assets i, j , excluding the pairs i, i . C_1 is the average of $C_{1t}(\tau)$ over time origins t .

The function C_{1t} is a measure of co-movement, from time t to time $t + \tau$, along the time t and the function C_1 is a measure of co-movement for the full period considered.

The interpretation of these functions is as follows: if the resulting function is close to 1, it means that there is no co-movement in the whole market; if it yields values lower than 1, then we can say that the stocks move together; and if, on the contrary, the values are greater than 1, the stocks tend to move in the opposite direction.

- C_2 . Consider the functions

$$C_{2t}(\tau) = \frac{\sum_{i,j} \delta x_i(t, \tau) \delta x_j(t, \tau)}{\sum_{i,j} (\delta x_i(t, \tau)^2 + \delta x_j(t, \tau)^2)}$$

$$C_2(\tau) = \langle C_{2t}(\tau) \rangle$$

where the same notation is used.

In this case, if the functions are close to 0 it means that there is no co-movement; if they are greater than 0, it means that the stocks move in the same direction and, when they are less than 0, the stocks move in the opposite direction.

- C_3 . Consider the functions

$$C_{3t}(\tau) = \frac{\sum_{i,j} \delta x_i(t, \tau) \delta x_j(t, \tau)}{\sum_{i,j} |\delta x_i(t, \tau) \delta x_j(t, \tau)|}$$

$$C_3(\tau) = \langle C_{3t}(\tau) \rangle$$

where the same notation is used.

These functions are interpreted in the same way that C_2 . Note that C_{3t} and C_3 take values between -1 and 1 .

Note that functions C_1 and C_2 are analogous to functions $\gamma(\tau)/\langle (\delta x)^2 \rangle$ and $C(\tau)$ in Puertas et al. [45], respectively, while function C_3 is a variation of C_2 . The three functions are intended to measure the degree of co-movement among all the stocks in the market, though they are slightly different.

3. Results

In this work, more than three thousand stocks (attending to liquidity, capitalization and free float criteria) in the USA market are considered from 2003 to March 2020. We will first show the results of the three functions proposed above, and the importance of subtracting the market to the evolution of the asset prices. Next, we will show how these parameters evolve in time and can be used to identify high-volatility periods. In the final subsection, other forms to represent the market are studied.

3.1. Co-Movement for Different Lag Times with and without Considering the Market

In this section, we want to study the behavior of the USA market through functions that show the co-movement of the shares. To calculate the shares, we will use, as data, the logarithm of the share price in two specific years, 2008 and 2018, in order to be able to check how the co-movement is affected in crisis stages and other more stable ones.

We will calculate the functions of the co-movement in two different ways: the first as described in the previous section, and the second by subtracting the market average, which is, $\delta x_i(t, \tau) = x_i(t + \tau) - x_i(t) - (m(t + \tau) - m(t))$, where x_i is the log price of asset i and $m = \langle x_i \rangle$ is the market (the mean of the log-price of all assets). The purpose is to check how the market average affects the stock co-movement calculations. Note that $m(t + \tau) - m(t) = \langle x_i(t + \tau) - x_i(t) \rangle$, which is, the market log-return is the equally weighted mean of the log-return of all the stocks.

The subtraction of the market is equivalent to removing the motion of the center of mass in physical systems. There, the dynamics of a system of particles can be separated in the motion of the center of mass, and the dynamics respect to it. Whereas external forces affect the center of mass, the internal dynamics is controlled by the internal forces.

We will use the three functions described in the previous section to measure the degree of co-movement.

In Figures 1–3, we show the results that were obtained with C_1 , C_2 , and C_3 , respectively, for the years 2008 and 2018, depending on whether or not the market average is taken into account.

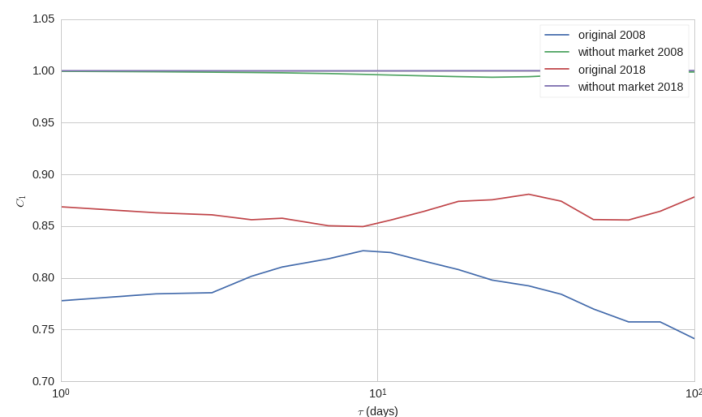


Figure 1. Co-movement with C_1 function. A value of 1 means that there is no co-movement, while values of less than 1 mean that the stocks tend to move in the same direction.

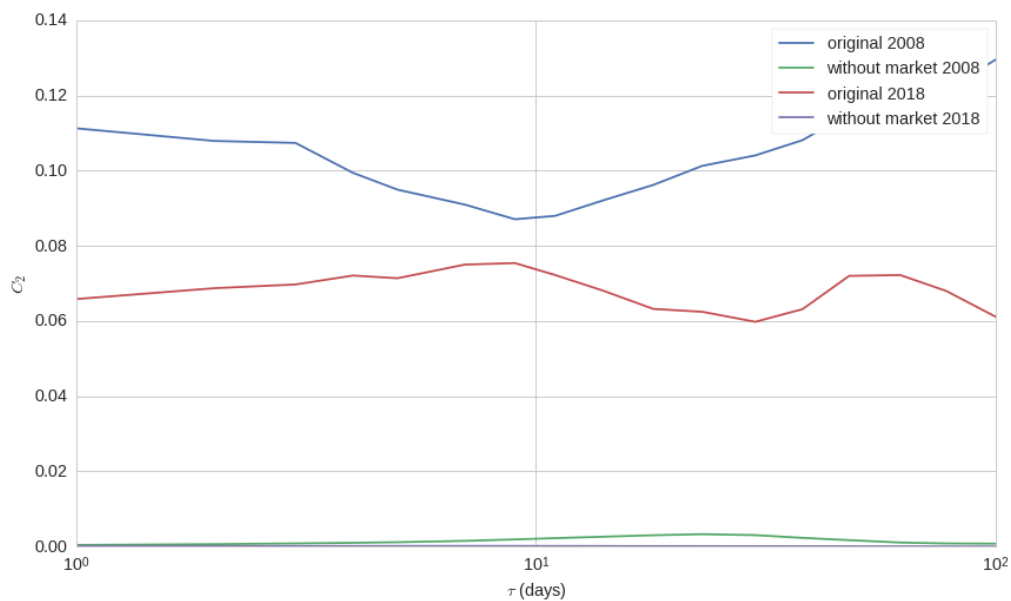


Figure 2. Co-movement with C_2 function. A value of 0 means that there is no co-movement, while values greater than 0 mean that the stocks tend to move in the same direction.

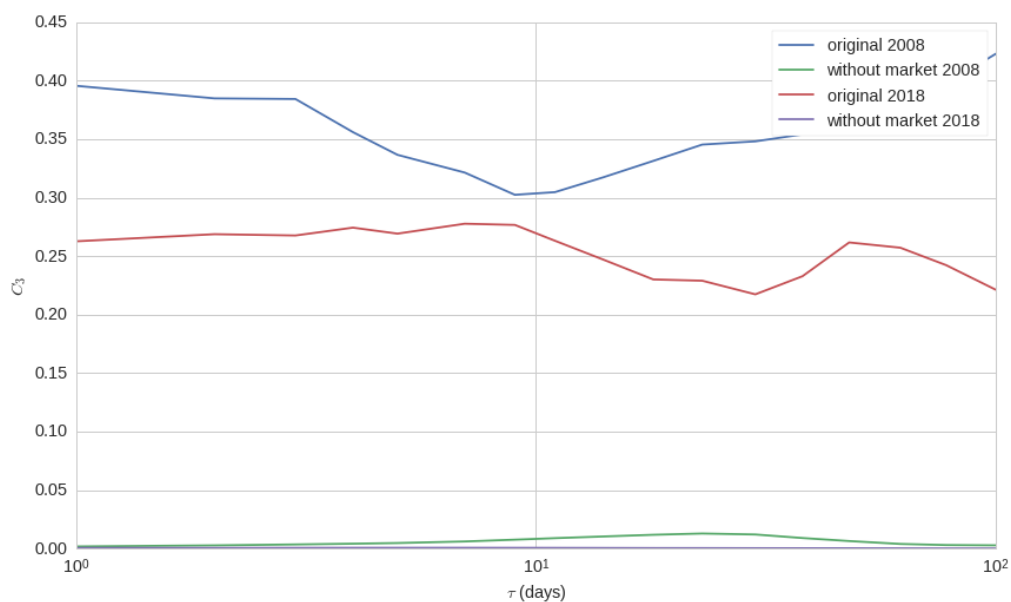


Figure 3. Co-movement with C_3 function. A value of 0 means that there is no co-movement, while values that are greater than 0 mean that the stocks tend to move in the same direction.

In both years and for the three functions, it is clear that there is some co-movement in the whole market (in the three cases, the co-movement in the year 2008 is a bit greater than in the year 2018), but all of the co-movement is explained by the market, since the co-movement disappears when we subtract the market.

Because the three functions essentially provide the same information, we use in the following $C_{3t}(\tau)$ and $C_3(\tau)$.

3.2. Co-Movement Along the Time

To study the evolution of the co-movement with a wide time horizon, we have used the function $C_{3t}(20)$, which is, the degree of co-movement after 20 working days.

In Figures 4 and 5, we show the function $C_{3t}(20)$ for the period (2003–2020). A moderate degree of co-movement can be observed at all times. Note that the co-movement is higher during crisis moments or bear markets: 2003 with the end of the dot.com bubble, the end of 2008 and the beginning of 2009 with the financial crisis, mid 2010 with the flash crash, mid 2011 with the Eurozone debt crisis, mid 2015 to early 2016 with the Chinese stock market crash, the end of 2018 with the close to 20% decline on major indexes and mid 2020 with the coronavirus crisis. At some particular days, the co-movement is very close to 1, which is an astonishing value of co-movement. Similar results have been reported by Tseng and Li [53] and Trinidad et al. [54].

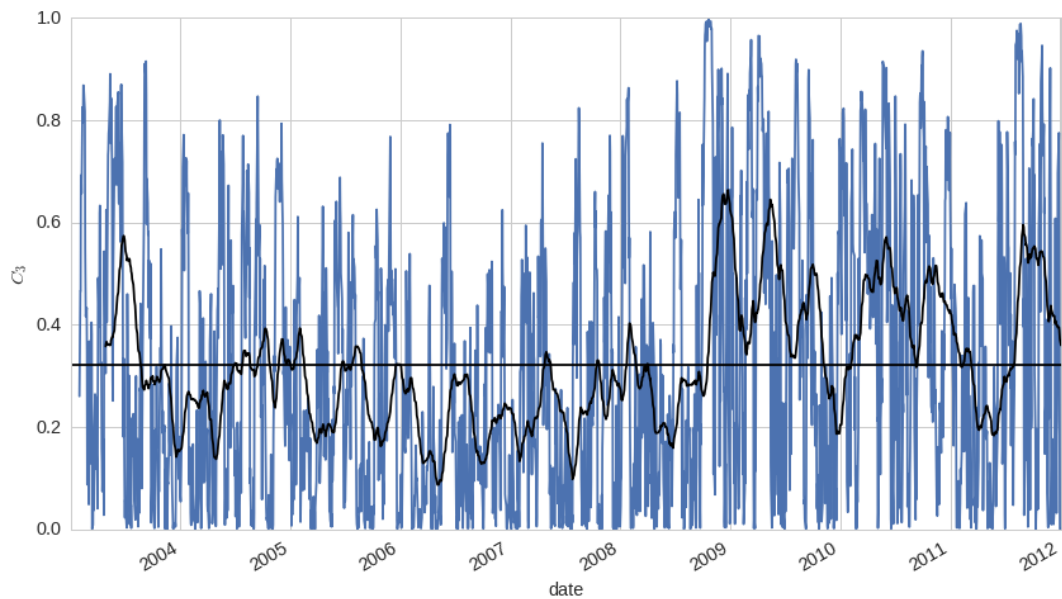


Figure 4. Co-movement ($C_{3t}(20)$) of the whole market from 2003 to 2011. Blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days.

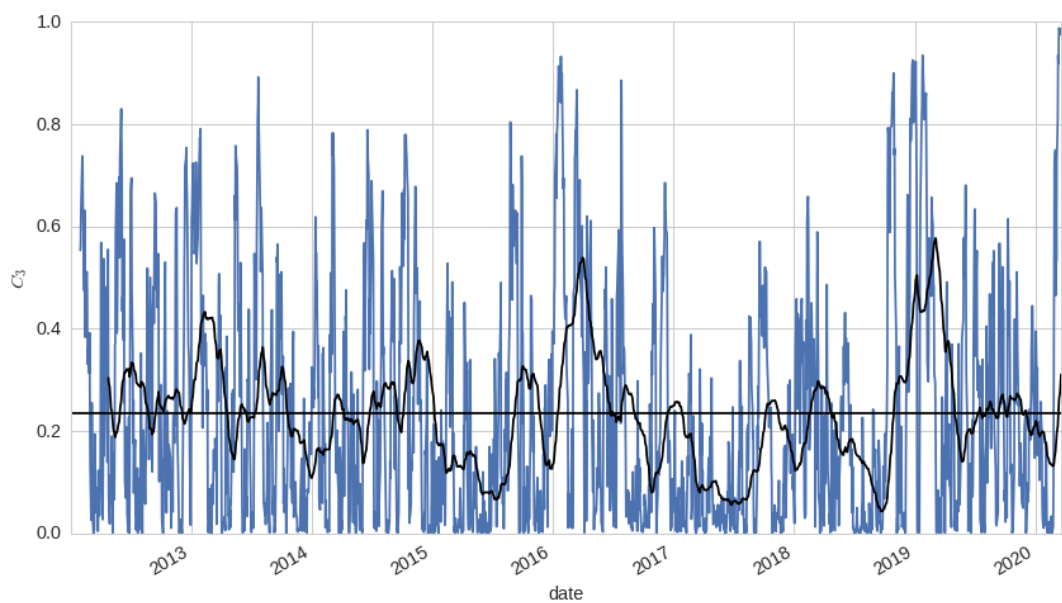


Figure 5. Co-movement ($C_{3t}(20)$) of the whole market from 2012 to 2020. Blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days.

In Figures 6 and 7, we show $C_{3t}(20)$ with the market subtracted, as explained in the previous section. Clearly, no co-movement is detected with this observable. The highest punctual peaks are less than 0.005.

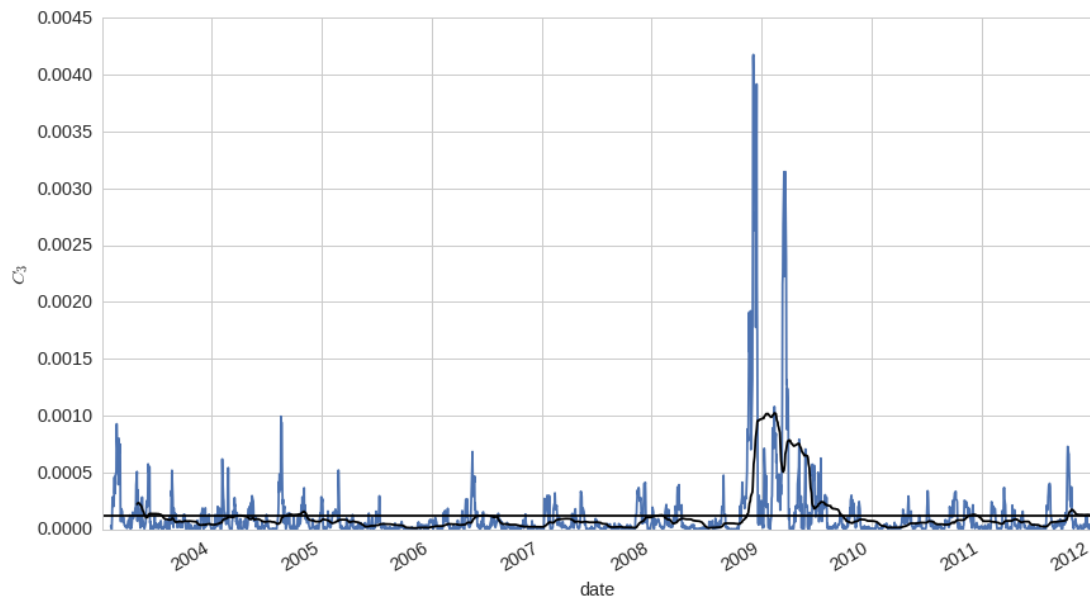


Figure 6. Co-movement ($C_{3t}(20)$) of the whole market from 2003 to 2011 with the market removed. Blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days

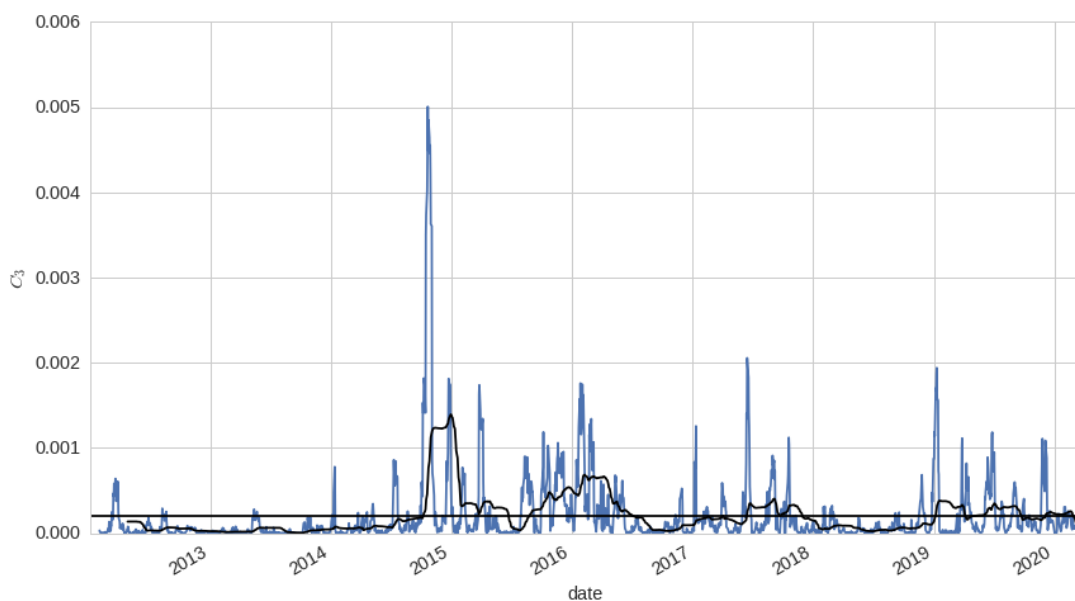


Figure 7. Co-movement ($C_{3t}(20)$) of the whole market from 2012 to 2020 with the market removed. Blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days

We can conclude that the market completely explains the co-movement among the stocks, which is quite a surprising result, since we have not considered any *beta* relationship.

3.3. Other Ways to Represent the Market

In this section, we have made the calculations taking into account different ways of representing the market, to see in which scenarios the co-movement is more reduced. Therefore, $\delta x_i(t, \tau)$ is calculated as $\delta x_i(t, \tau) = x_i(t + \tau) - x_i(t) - (m(t + \tau) - m(t))$, where x_i is the log price of asset i and $m(t)$ represent the market. The ways of studying the market that we have selected are:

- ew: equal weight. This is the representation described in the previous section, which is, $m(t) = \frac{1}{N(t)} \sum_i x_i(t)$, where $N(t)$ is the number of stocks at time t .
- cap: this representation is calculated as a capitalization-weighted average, which is, $m = \sum_j w_j x_j$ and $w_j = c_j / \sum_k c_k$ with c_j the capitalization of asset j .
- SPY: the SP500 index.
- IWM: the Russell 2000 index.

In the previous section, we studied the co-movement in two periods 2003–2011 and 2012–2020, and found that, if we measure the market with the ew method, the co-movement in these two periods is very close to zero. In Scheme 1, it is shown what happens when we consider other ways to represent the market. It can be seen that the representations of the market that explain the co-movement better than any of the alternatives considered are ew.

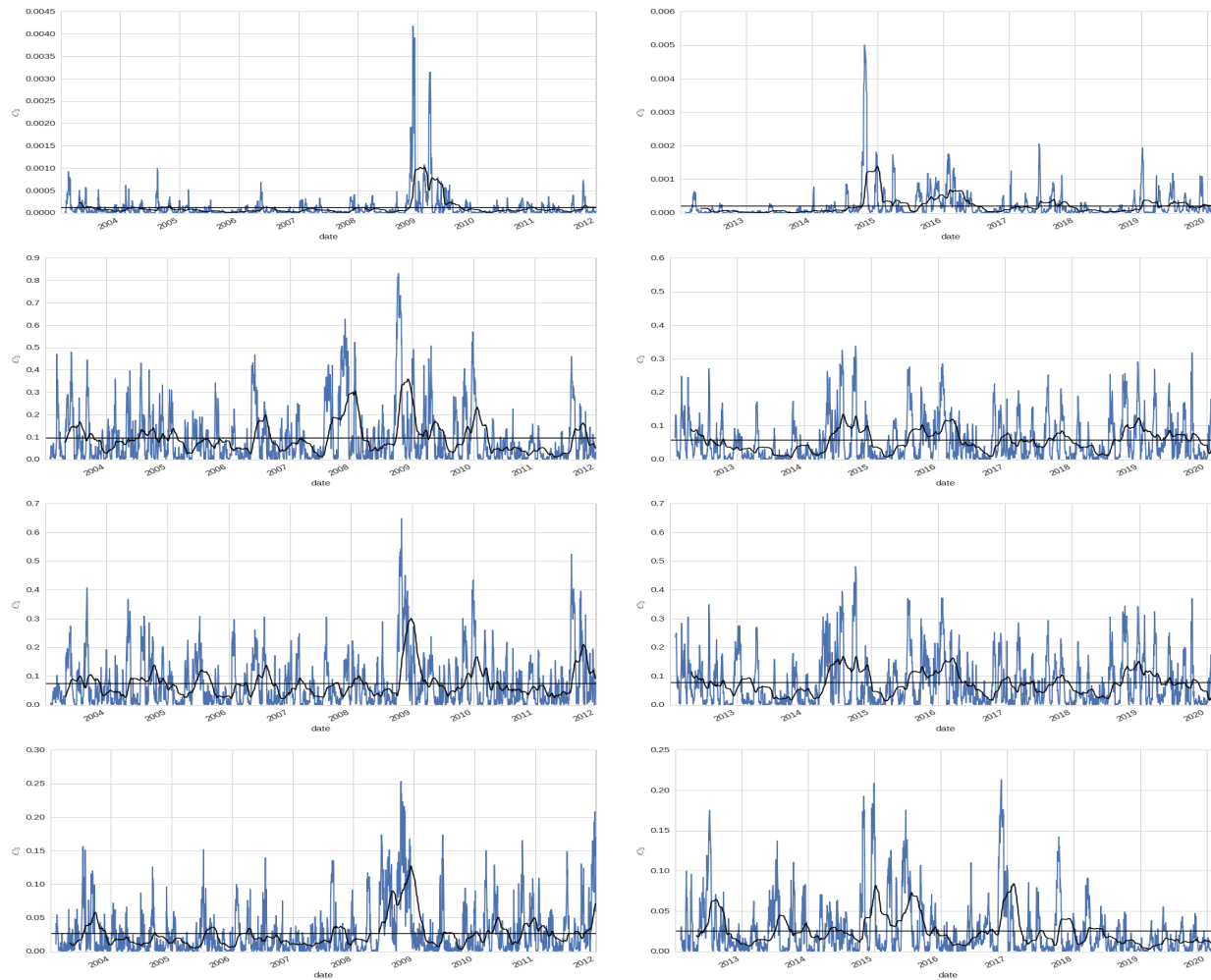
3.3.1. Considering Beta

As we commented in the introduction, the market beta is used in the literature to model the log-return of a stock with respect to the log-return of the market. The simplest model is the Sharpe one, but other models, like the Fama–French factor model, also consider the market beta as a part of the model. Subsequently, the log-return of a stock is modeled as $x_i(t + \tau) - x_i(t) = \alpha_i + \beta_i(m(t + \tau) - m(t)) + \varepsilon_i$, where ε_i follows a fixed stationary distribution with zero mean and is independent of the market. It follows that $\frac{x_i(t+\tau) - x_i(t)}{\beta_i} - (m(t + \tau) - m(t)) = \frac{\alpha_i}{\beta_i} + \frac{\varepsilon_i}{\beta_i}$, and hence it is independent of the market.

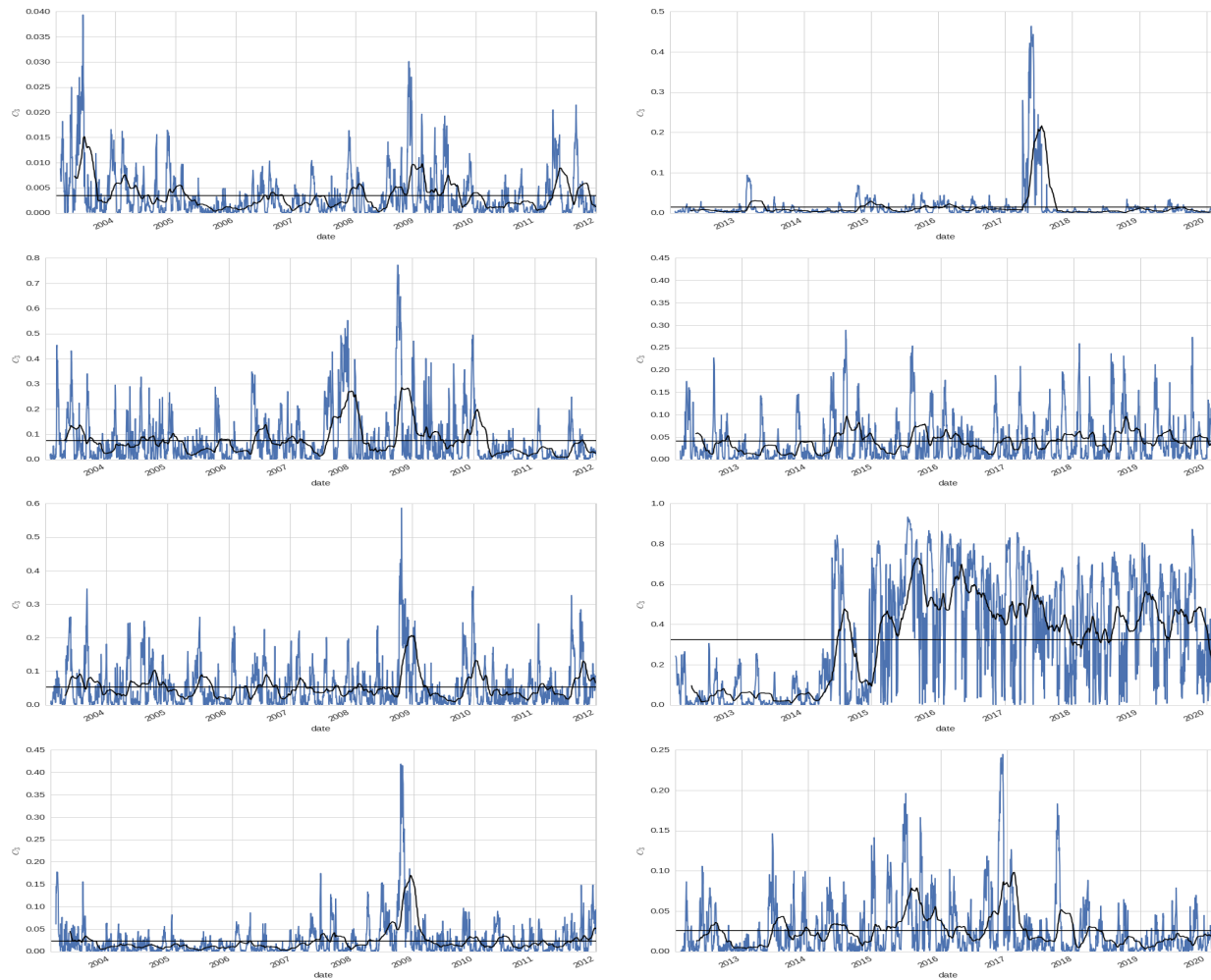
To check this model, in this case, $\delta x_i(t, \tau)$ is calculated as $\delta x_i(t, \tau) = \frac{x_i(t+\tau) - x_i(t)}{\beta_i} - (m(t + \tau) - m(t))$, and m is represented in the different ways described previously (ew, cap, SPY, IWM). Assuming the beta model, the co-movement that is calculated in this way should be lower than the co-movement calculated without considering the beta.

In Scheme 2, we show the co-movement ($C_{3i}(20)$) calculated within the Fama and French model. We can say that this version does not improve the results obtained previously and hence the previous model is preferred, since the results are better and the model is simpler.

Therefore, from the point of view of co-movement, we can conclude that the relationship between the log-returns of a stock and the market is provided by $x_i(t + \tau) - x_i(t) = \alpha_i + (m(t + \tau) - m(t)) + \varepsilon_i$, which implies $\beta = 1$.



Scheme 1. Co-movement ($C_{3f}(20)$) of the whole market during 2003-2011 (left) and 2012-2020 (right) with the market removed. Blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days. The representation of the market is, from top to bottom: ew, cap, SPY, IWM. The horizontal line represent the average co-movement ($C_3(20)$) along the full period.



Scheme 2. Co-movement ($C_{3t}(20)$) of the whole market during 2003–2011 (left) and 2012–2020 (right) with the market removed when considering the beta of each stock. The blue line is the daily co-movement, while black line is a moving average of the blue line with a time window of 60 trading days. The representation of the market is, from top to bottom: ew, cap, SPY, IWM. The horizontal line represent the average co-movement ($C_3(20)$) along the full period.

4. Conclusions

In this paper, a new approach to study co-movement among the stocks of the whole (USA) market is considered, by using some functions that are inspired by cooperative dynamics in physical particle systems. We prove that these functions identify the economic crisis as periods of increased co-movement, which is agreement with the finding of previous research (see, for example, Tseng and Li [53] and Trinidad et al. [54]). From this perspective, the market (represented by an equally weighted portfolio of all the stocks considered) completely explains all of the co-movement among the stocks in such a way that, if we remove the market, the co-movement is almost zero. This fact supposes an important finding considering that one of the most important inconveniences of beta estimations is their instability in time.

The three functions are three different ways to measure the whole co-movement of the market. The fact that the results are similar implies that the measure is quite robust. Therefore, these measures seem to be quite useful, since they provide a way to measure the co-movement among the stocks and we can obtain results that are not obvious. Moreover, we think that these measures open a new door to study the co-movement among a set of stocks for different purposes.

Other ways to represent the market (capitalization weighted, SP500 index and Russell 2000 index) can also explain the co-movement in the market, but to a lesser extent than an equally weighted portfolio. This result is similar to a recent finding of Lopez et al. [20], where the authors show that a market factor represented by an equally weighted portfolio is highly significant in a factor model.

To conclude, the market is also considered taking into account the market beta of the stocks, but again, the results are not as good as the simpler equally weighted portfolio. These results lead us to propose a modification of the CAPM model, removing the beta in the classic model.

Author Contributions: All authors contributed to conceptualization, methodology, writing and reviewing. MNLG and JETS contributed to the financial aspects of the paper, while MASG, AMP and FJDIN contributed to the mathematical and physical models. All authors have read and agreed to the published version of the manuscript.

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References

- Sharpe, W. Capital Asset Prices: A theory of market equilibrium under conditions of risk. *J. Financ.* **1964**, *19*, 425–442.
- Lintner, J. The valuation of risk assets and the selection of risky investments in stock portfolios and capital budgets. *Rev. Econ. Stat.* **1965**, *47*, 13–37.
- Black, F. Capital Market Equilibrium with Restricted Borrowing. *J. Bus.* **1972**, *45*, 444–455.
- Black, F. Beta and returns. *J. Portf. Manag. Fall 1993*, 8–18.
- Blume, M. On the assessment of risk. *J. Financ.* **1971**, *26*, 1–10.
- Black, F.; Jensen, M.; Scholes, M. The capital asset pricing model: Some empirical tests. *Stud. Theory Cap. Mark.* **1972**, *81*, 79–121.
- Fama, E.F.; MacBeth, J.D. Risk, Return, and Equilibrium: Empirical Tests. *J. Political Econ.* **1973**, *81*, 607–636.
- King, B.F. Market and Industry Factors in Stock Price Behavior. *J. Bus.* **1966**, *39*, 139–190.
- Meyers, S.L. A Re-Examination of Market and Industry Factors in Stock Price Behavior. *J. Financ.* **1973**, *28*, 695–705.
- Banz, R.W. The relationship between return and market value of common stocks. *J. Financ. Econ.* **1981**, *9*, 3–18.
- Roll, R. R^2 . *J. Financ.* **1988**, *43*, 541–566.
- Bhandari, L.C. Debt/Equity Ratio and Expected Common Stock Returns: Empirical Evidence. *J. Financ.* **1988**, *43*, 507–528.

13. Stattman, D. Book values and stocks returns. *Chic. MBA J. Sel. Pap.* **1980**, *4*, 25–45.
14. Rosenberg, B.; Reid, K.; Lanstein, R. Persuasive Evidence of Market Inefficiency. *J. Portf. Manag.* **1985**, *11*, 9–17.
15. Fama, E.F.; French, K.R. The Cross-Section of Expected Stock Returns. *J. Financ.* **1992**, *47*, 427–465.
16. Isakov, D. Is beta still alive? Conclusive evidence from the Swiss stock market. *Eur. J. Financ.* **1999**, *5*, 202–212.
17. Pettengill, G.N.; Sundaram, S.; Mathur, I. The Conditional Relation between Beta and Returns. *J. Financ. Quant. Anal.* **1995**, *30*, 101–116.
18. Chan, L.; Lakonishok, J. Institutional trades and intraday stock price behavior. *J. Financ. Econ.* **1993**, *33*, 173–199.
19. Grundy, K.; Malkiel, B.G. Reports of beta's death have been greatly exaggerated. *J. Portf. Manag.* **1996**, *22*, 36–44.
20. López-García, M.N.; Trinidad-Segovia, J.E.; Sánchez-Granero, M.A.; Pouchkarev, I. Extending the Fama and French model with a long term memory factor. *Eur. J. Oper. Res.* **2020**, in press, doi:10.1016/j.ejor.2019.07.071.
21. Domowitz, I.; Hansch, O.; Wang, X. Liquidity commonality and return co-movement. *J. Financ. Mark.* **2005**, *8*, 351–376.
22. Byrne, J.P.; Cao, S.; Korobilis, D. Decomposing global yield curve co-movement. *J. Bank. Financ.* **2019**, *106*, 500–513.
23. Morck, R.; Yeung, B.; Yu, W. The information content of stock markets: Why do emerging markets have synchronous stock price movements? *J. Financ. Econ.* **2000**, *58*, 215–260.
24. Jach, A. International Stock market Comovement Time Scale Outlined A Thick Pen. *J. Empir. Financ.* **2017**, *43*, 115–129.
25. Parsley, D.; Popper, H. Return comovement. *J. Bank. Financ.* **2020**, *112*, 105–223.
26. Bonfiglioli, A.; Favero, C.A. Explaining co-movements between stock markets: The case of US and Germany. *J. Int. Money Financ.* **2005**, *24*, 1299–1316.
27. Rua, A.; Nunes, L.C. International comovement of stock market returns: A wavelet analysis. *J. Empir. Financ.* **2009**, *16*, 632–639.
28. Akoum, I.; Graham, M.; Kivihaho, J.; Nikkinen, J.; Omran, M. Co-movement of oil and stock prices in the GCC region: A wavelet analysis. *Q. Rev. Econ. Financ.* **2012**, *52*, 385–394.
29. Reboredo, J.C. Modelling oil prices and exchange rates co-movements. *J. Policy Model.* **2012**, *34*, 419–440.
30. Magdaleno, M.; Pinho, C. International stock market indices comovements: A new look. *Int. J. Financ. Econ.* **2012**, *17*, 89–102.
31. Loh, L. Co-movement of Asia-Pacific with European and US stock market returns: A cross-time-frequency analysis. *Res. Int. Bus. Financ.* **2013**, *29*, 1–13.
32. Baca, S.; Garbe, B.; Weiss, R. The rise of sector effects in major equity markets. *Financ. Anal. J.* **2000**, *56*, 35–40.
33. Cavaglia, S.; Brightman, C.; Aked, M. The increasing importance of industry factors. *Financ. Anal. J.* **2000**, *56*, 41–54.
34. Griffin, J.; Karolyi, A. Another look at the role of industrial structure of markets for international diversification strategies. *J. Financ. Econ.* **1998**, *50*, 351–373.
35. L'Her, J.; Sy, O.; Tnani, Y. Country, industry and risk factor loadings in portfolio management. *J. Portf. Manag.* **2002**, *28*, 70–79.
36. Brooks, R.; Del Negro, M. Firm-Level Evidence on International Stock Market Comovement. *Rev. Financ.* **2006**, *10*, 69–98.
37. Antonakakis, N.; Chatziantoniou, I.; Filis, G. Dynamic Co-movements of Stock Market Returns, Implied Volatility and Policy Uncertainty. *Econ. Lett.* **2013**, *120*, 87–92.
38. Fernández-Avilés Calderón, G.; Montero, J.M.; Orlov, A. Spatial Modeling of Stock Market Comovements. *Financ. Res. Lett.* **2012**, *9*, 202–212.
39. Capiello, L.; Gerard, B.; Kadareja, A.; Manganelli, S. Measuring Comovements by Regression Quantiles. *J. Financ. Econom.* **2014**, *12*, 645–678.
40. Garcia, R.; Tsafack, G. Dependence structure and extreme comovements in international equity and bond markets. *J. Bank. Financ.* **2011**, *35*, 1954–1970.
41. Fernández-Avilés, G.; Montero, J.M.; Sanchis-Marco, L. Extreme downside risk co-movement in commodity markets during distress periods: A multidimensional scaling approach. *Eur. J. Financ.* **2020**, *12*, 1207–1237.

42. Ramos Requena, J.P.; Trinidad Segovia, J.E.; Sánchez Granero, M.A. An Alternative Approach to Measure Co-Movement between Two Time Series. *Mathematics* **2020**, *8*, 261, doi:10.3390/math8020261.
43. Clara-Rahola, J.; Puertas, A.M.; Sánchez-Granero M.A.; Trinidad-Segovia, J.E.; de las Nieves, F.J. Diffusive and arrestedlike dynamics in currency exchange markets. *Phys. Rev. Lett.* **2017**, *118*, 068301.
44. Sánchez-Granero M.A.; Trinidad-Segovia, J.E.; Clara-Rahola, J.; Puertas, A.M.; de las Nieves, F.J. A model for foreign exchange markets based on glassy Brownian systems. *PLoS ONE* **2017**, *12*, e0188814.
45. Puertas, A.M.; Sánchez-Granero, M.A.; Clara-Rahola, J.; Trinidad-Segovia, J. E.; de las Nieves, F.J. Stock markets: A view from soft matter. *Phys. Rev. E* **2020**, *101*, 032307.
46. Donati, C.; Glotzer, S.C.; Poole, P.H. Growing spatial correlations of particle displacements in a simulated liquid on cooling toward the glass transition. *Phys. Rev. Lett.* **1999**, *82*, 5064.
47. Cates, M.E.; Evans, M.R. *Soft and Fragile Matter, Nonequilibrium Dynamics, Metastability and Flow*; Institute of Physics: Bristol, UK, 2000.
48. Weeks, E.; Crocker, J.C.; Levitt, A.C.; Schofield, A.; Weitz, D.A. Three-dimensional direct imaging of structural relaxation near the colloidal glass transition. *Science* **2000**, *287*, 627.
49. Glotzer, S.C.; Novikov, V.N.; Schroeder, T.B. Time dependent, four-point density correlation function description of dynamical heterogeneity and decoupling in supercooled liquids. *J. Chem. Phys.* **2000**, *112*, 509–512.
50. Berthier, L.; Biroli, G.; Bouchad, J.-P.; Cipelletti, L.; van Saarloos, W. *Dynamical Heterogeneities in Glasses, Colloids and Granular Media*; Oxford University Press: Oxford, UK, 2011.
51. Muranaka, T.; Hiwatari, Y. β relaxation in a highly supercooled state via molecular dynamics simulation. *Phys. Rev. E* **1995**, *51*, R2735(R).
52. Zahn, K.; Lenke, R.; Maret, G. Two-stage melting of paramagnetic colloidal crystals in two dimensions. *Phys. Rev. Lett.* **1999**, *82*, 2721.
53. Tseng, J.J.; Li, S.P. Asset returns and volatility clustering in financial time series. *Phys. A Stat. Mech. Appl.* **2011**, *390*, 1300–1314.
54. Trinidad Segovia, J.E.; Fernández-Martínez, M.; Sánchez-Granero, M.A. A novel approach to detect volatility clusters in financial time series. *Phys. A Stat. Mech. Appl.* **2019**, *535*, 122452.



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