Almost Optimal Testers for Concise Representations

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— Abstract

We give improved and almost optimal testers for several classes of Boolean functions on n variables that have concise representation in the uniform and distribution-free model. Classes, such as k-Junta, k-Linear, s-Term DNF, s-Term Monotone DNF, r-DNF, Decision List, r-Decision List, size-s Decision Tree, size-s Boolean Formula, size-s Branching Program, s-Sparse Polynomial over the binary field and functions with Fourier Degree at most d.

The approach is new and combines ideas from Diakonikolas et al. [24], Bshouty [13], Goldreich et al. [32], and learning theory. The method can be extended to several other classes of functions over any domain that can be approximated by functions with a small number of relevant variables.

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1 Introduction

Property testing of Boolean function was first considered in the seminal works of Blum, Luby and Rubinfeld [12] and Rubinfeld and Sudan [46] and has recently become a very active research area. See for example, [2, 4, 5, 6, 8, 9, 13, 15, 16, 17, 18, 19, 20, 21, 24, 27, 31, 33, 36, 37, 41, 40, 43, 47] and other works referenced in the surveys and books [29, 30, 44, 45].

A Boolean function $f : \{0, 1\}^n \to \{0, 1\}$ is said to be k-junta if it depends on at most k coordinates. The class k-Junta is the class of all k-juntas. The class k-Junta has been of particular interest to the computational learning theory community [10, 11, 14, 23, 34, 38, 42]. A problem closely related to learning k-Junta is the problem of learning and testing subclasses C of k-Junta and classes C of Boolean functions that can be approximated by k-juntas [9, 11, 25, 17, 24, 32, 33, 43]. In both testing and learning we are given black-box query access to a Boolean function f. In learning, for $f \in C$, we need to learn, with high probability, a hypothesis h that is ϵ -close to f. In testing, for any Boolean function f, we need to distinguish, with high probability, the case that f is in C versus the case that f is ϵ -far from every function in C.

In the uniform-distribution property testing (and learning) model, the distance between Boolean functions is measured with respect to the uniform distribution. In the distributionfree property testing, [32], (and learning [48]) the distance between Boolean functions is measured with respect to an arbitrary and unknown distribution \mathcal{D} over $\{0, 1\}^n$. In the





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5:2 Optimal Testers

distribution-free model, the testing (and learning) algorithm is allowed (in addition to making black-box queries) to draw random $x \in \{0, 1\}^n$ according to the distribution \mathcal{D} . This model is studied in [13, 22, 26, 28, 35, 39].

1.1 Results

In Table 1, we list all the previous results and our results in this paper. In the table, O(T) stands for $O(T \cdot Poly(\log T))$, U and D stand for uniform and distribution-free models, resp., and Exp and Poly stand for exponential and polynomial time, resp.

It follows from the lower bounds of Saglam [47], that our query complexity is almost optimal (up-to log-factor) for the classes k-Junta, k-Linear, k-Term, s-Term DNF, s-Term Monotone DNF, r-DNF (r constant), Decision List, r-Decision List (r constant), size-s Decision Tree, size-s Branching Programs and size-s Boolean Formula.

1.2 Notations

In this subsection, we give some notations that we use throughout the paper.

Denote $[n] = \{1, 2, \ldots, n\}$. For $S \subseteq [n]$ and $x = (x_1, \ldots, x_n)$ we denote $x(S) = \{x_i | i \in S\}$. For $X \subset [n]$ we denote by $\{0, 1\}^X$ the set of all binary strings of length |X| with coordinates indexed by $i \in X$. For $x \in \{0, 1\}^n$ and $X \subseteq [n]$ we write $x_X \in \{0, 1\}^X$ to denote the projection of x over coordinates in X. We denote by 1_X and 0_X the all-one and all-zero strings in $\{0, 1\}^X$, respectively. When y is a variable then $z = (y)_X$ is the all y string with coordinates indexed by $i \in X$, i.e., $z_i = y$ for all $i \in X$. When we write $x_I = 0$ we mean $x_I = 0_I$. For $X_1, X_2 \subseteq [n]$ where $X_1 \cap X_2 = \emptyset$ and $x \in \{0, 1\}^{X_1}, y \in \{0, 1\}^{X_2}$ we write $x \circ y$ to denote their concatenation, i.e., the string in $\{0, 1\}^{X_1 \cup X_2}$ that agrees with x over coordinates in X_1 and agrees with y over coordinates in X_2 . For $X \subseteq [n]$ we denote $\overline{X} = [n] \setminus X = \{x \in [n] | x \notin X\}$. For a function $f : \{0, 1\}^k \to \{0, 1\}, x \in \{0, 1\}^n$ and $X = \{i_1, \ldots, i_k\} \subseteq [n]$ where $i_1 < i_2 < \cdots < i_k$, when we write $f(x_X)$ we mean $f(x_{i_1}, \ldots, x_{i_k})$.

Given $f, g: \{0, 1\}^n \to \{0, 1\}$ and a probability distribution \mathcal{D} over $\{0, 1\}^n$, we say that f is ϵ -close to g with respect to \mathcal{D} if $\mathbf{Pr}_{x\in\mathcal{D}}[f(x)\neq g(x)]\leq\epsilon$, where $x\in\mathcal{D}$ means x is chosen from $\{0, 1\}^n$ according to the distribution \mathcal{D} . We say that f is ϵ -far from g with respect to \mathcal{D} if $\mathbf{Pr}_{x\in\mathcal{D}}[f(x)\neq g(x)]\geq\epsilon$. For a class of Boolean functions C, we say that f is ϵ -far from g with respect to \mathcal{D} . We will use U to denote the uniform distribution over $\{0, 1\}^n$ or over $\{0, 1\}^X$ when X in clear from the context.

For a distribution \mathcal{D} over $\{0,1\}^n$ and $X \subset [n]$, we denote by \mathcal{D}_X the distribution \mathcal{D} projected on the coordinates X. That is, the distribution of x_X when $x \in \mathcal{D}$.

For a Boolean function f and $X \subset [n]$, we say that X is an *influential set* of f if there are $a, b \in \{0, 1\}^n$ such that $f(a) \neq f(b_X \circ a_{\overline{X}})$. We call the pair (a, b) (or just a when b = 0) a witness of f for the influential set X. When $X = \{i\}$ then we say that x_i is an *influential* variable of f and a is a witness of f for x_i . Obviously, if X is influential set of f then x(X)contains at least one influential variable of f.

We say that the Boolean function $f : \{0,1\}^n \to \{0,1\}$ is a *literal (dictatorship* and *anti-dictatorship*, resp.) if $f \in \{x_1, \ldots, x_n, \overline{x_1}, \ldots, \overline{x_n}\}$ where \overline{x} is the negation of x ($f \in \{x_1, \ldots, x_n\}$ and $f \in \{\overline{x_1}, \ldots, \overline{x_n}\}$, resp.).

Let C be a class of Boolean functions $f : \{0,1\}^n \to \{0,1\}$. We say that C is symmetric if for every permutation $\pi : [n] \to [n]$ and every $f \in C$ we have $f_{\pi} \in C$ where $f_{\pi}(x) := f(x_{\pi(1)}, \dots, x_{\pi(n)})$. We say that C is closed under zero projection (resp. closed under one projection) if for every $f \in C$ and every $i \in [n]$, $f(0_{\{i\}} \circ x_{\overline{\{i\}}}) \in C$ (resp. $f(1_{\{i\}} \circ x_{\overline{\{i\}}}) \in C$). We say it is closed under zero-one projection if is closed under zero and one projection.

Table 1 A table of the results. In the table, $\tilde{O}(T)$ stands for $O(T \cdot poly(\log T))$, U and D stand for
uniform and distribution-free models, resp., and Exp and Poly stand for exponential and polynomial
time, resp.

Class of Functions	Model	$\# \mathbf{Queries}$	Time	Reference
s-Term Monotone DNF	U	$ ilde{O}(s^2/\epsilon)$	Poly.	[43]
s-Term Unate DNF	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
	U	$ ilde{O}(s/\epsilon)$	Poly.	This Paper
s-Term Monotone r -DNF	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
s-Term Unate r -DNF	U	$ ilde{O}(s/\epsilon)$	Poly.	This Paper
	D	$ ilde{O}(s^2 r/\epsilon)$	Poly.	This Paper
s-Term DNF	U	$ ilde{O}(s^2/\epsilon)$	Exp.	[24]
	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
	U	$ ilde{O}(s/\epsilon)$	Exp.	This Paper
r-DNF (Constant r)	U	$\tilde{O}(1/\epsilon)$	Poly.	This Paper
Decision List	U	$\tilde{O}(1/\epsilon^2)$	Poly.	[24]
	U	$\tilde{O}(1/\epsilon)$	Poly.	This Paper
Length- k Decision List	D	$ ilde{O}(k^2/\epsilon)$	Poly.	This Paper
r-DL (Constant r)	U	$\tilde{O}(1/\epsilon)$	Poly.	This Paper
k-Linear	U	$ ilde{O}(k/\epsilon)$	Poly.	[7, 12]
	D	$ ilde{O}(k/\epsilon)$	Poly.	This Paper
k-Term	U	$O(1/\epsilon)$	Poly.	[43]
	U	$\tilde{O}(1/\epsilon)$	Poly.	This Paper
	D	$ ilde{O}(k/\epsilon)$	Poly.	This Paper
size- s Decision Trees and	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
size-s Branching Programs	U	$ ilde{O}(s/\epsilon)$	Exp.	This Paper
	D	$ ilde{O}(s^2/\epsilon)$	Exp.	This Paper
size-s Boolean Formulas	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
	U	$ ilde{O}(s/\epsilon)$	Exp.	This Paper
size- s Boolean Circuit	U	$ ilde{O}(s^2/\epsilon^2)$	Exp.	[17]
	U	$ ilde{O}(s^2/\epsilon)$	Exp.	This Paper
Functions with	U	$\tilde{O}(2^{2d}/\epsilon^2)$	Exp.	[17]
Fourier Degree $\leq d$	D	$\tilde{O}(2^d/\epsilon + 2^{2d})$	Poly.	This Paper
s-Sparse Polynomial	U	$poly(s/\epsilon) + \tilde{O}(2^{2d})$	Poly.	[1, 25]
over F_2 of Degree d	U	$\tilde{O}(s^2/\epsilon + 2^{2d})$	Poly.	This Paper+[1]
	U	$\tilde{O}(s/\epsilon + s2^d)$	Poly.	This Paper
	D	$\tilde{O}(s^2/\epsilon + s2^d)$	Poly.	This Paper
s-Sparse Polynomial	U	$ ilde{O}(s/\epsilon^2)$	Exp.	[17]
over F_2	U	$Poly(s/\epsilon)$	Poly.	[25]
	U	$\tilde{O}(s^2/\epsilon)$	Poly.	This Paper

2 Overview of the Distribution-Free Tester

2.1 Preface

Our approach refers to testing properties that are (symmetric) sub-classes C of k-juntas; that is, $f : \{0, 1\}^n \to \{0, 1\}$ has the property if there exists a function $f' : \{0, 1\}^k \to \{0, 1\}$ that belongs to a predetermined class C' of functions (over k-bit strings) such that $f(x) = f'(x_{\Gamma})$

for some k-subset Γ . Our new approach builds upon the "testing by implicit sampling" approach of Diakonikolas et al. [24], while extending it from the case of uniform distribution to the case of arbitrary unknown distributions (i.e., the distribution-free model).

This allows us to present (almost optimal) distribution-free testers for classes of properties that are sub-classes of k-juntas, which correspond to classes of k-bit long Boolean functions. While we follow Diakonikolas et al. [24] in considering learning algorithms for the under-

lying classes, our approach is also applicable to testing algorithms (see [30, Sec. 6.2]).

Let us again spell out our task. For a class C of n-bit long Boolean functions and a proximity parameter ϵ , given samples from an unknown distribution \mathcal{D} and oracle access to a function $f: \{0,1\}^n \to \{0,1\}$, we wish to distinguish the case that $f \in C$ from the case that f is ϵ -far from C. Recall that C is a (symmetric) class consisting of a symmetric subclass of k-juntas C'; that is, $f \in C$ if and only if there exists a k-subset $\Gamma \subset [n]$ and $f' \in C'$ such that $f(x) = f'(x_{\Gamma})$, where $x_{\{i_1,\ldots,i_k\}} = (x_{i_1},\ldots,x_{i_k})$. Actually, we also assume that C' is closed under zero projection.

2.2 A Bird's Eye View

The basic strategy is to consider a random partition of [n] to $r = O(k^2)$ parts, denoted (X_1, \ldots, X_r) , while relying on the fact that, whp, each X_i contains at most one influential variable. Assuming that $f \in C$, first we determine a set I of at most k indices such that $\bigcup_{i \in [n] \setminus I} X_i$ contains no "significantly influential" variables of f. Suppose that $f' : \{0, 1\}^k \to \{0, 1\}, f' \in C'$, is a function that corresponds to the tested function $f : \{0, 1\}^n \to \{0, 1\}$, and that $I \subset [n]$ is indeed the collection of all sets that contain influential variables. The crucial ingredient is devising a method that allows to generate samples of the form (x', f'(x')), when given samples of the form (x, f(x)) (for $x \in \mathcal{D}$). We stress that we cannot afford to find the influential variables, and so this method works without determining these locations. Using this method, we can test whether f' belongs to the underlying class C'; hence, we test f by implicitly sampling the projection of \mathcal{D} on the (unknown) influential variables.

The method employed by Diakonikolas et al. [24] only handles the uniform distribution (i.e., the case that \mathcal{D} is uniform over $\{0,1\}^n$), and so it only yields testers for the standard testing model (rather than for the distribution-free testing model). Furthermore, their method as well as the identification of the set I rely heavily on the notion of influence of sets, where the influence of a set S of locations on the value of a function is defined as $\Pr_{x',x''\in\{0,1\}^n:x'_S=x''_S}[f(x')\neq f(x'')]$. However, this notion refers to the uniform distribution (over $\{0,1\}^n$) and does not adequate for the distribution-free context (e.g., for $f(x) = x_1 + x_2$ we may get $\Pr_{x',x''\in\mathcal{D}:x'_1=x''_1}[f(x')\neq f(x'')] = 0$).

We use a different way of identifying the set I and for generating samples for the underlying function f'. Loosely speaking, we identifies I as the set of indices i for which $f(1_{X_i} \circ 0_{\overline{X_i}}) \neq f(0^n)$, where (recall that) $1_S \circ 0_{\overline{S}}$ is a string that is 1 on the locations in S and is 0 on other locations. (Be warned that this description is an over-simplification!) This means that for every $i \in I$ and $x \in \{0, 1\}^n$, the value of x at the influential variable in the set X_i (a variable whose location is unknown to us!), equals $f(x') + f(0^n)$ where $x' = x_{X_i} \circ 0_{\overline{X_i}}$, i.e., $x'_j = x_j$ if $j \in X_i$ and $x'_j = 0$ otherwise.² Note that the foregoing holds when $f \in C$; in

$$f(x') = x_{\tau(i)} \cdot f(1_{X_i} \circ 0_{\overline{X_i}}) + (x_{\tau(i)} + 1) \cdot f(0^n) = x_{\tau(i)} + f(0^n)$$

since $f(1_{X_i} \circ 0_{\overline{X_i}}) + f(0^n) = 1.$

¹ The addition operation in this paper is over the binary field F_2 .

² Indeed, if $\tau(i) \in X_i$ is the index of the (unique) influential variable that resides in the set X_i , then

general, we can test whether $x \mapsto f(x') + f(0^n)$ is close to a dictatorship (under the uniform distribution) and reject otherwise, whereas if the mapping is close to a dictatorship, we can self-correct it.

To sample the distribution \mathcal{D}_{Γ} , where Γ is the influential variables in $X_I = \bigcup_{i \in I} X_i$, we sample \mathcal{D} and determine the value of the influential variable in each set X_i , for $i \in I$. Queries to the function f' are answered by querying f such that the query $y = y_1 \cdots y_k$ is mapped to the query $\operatorname{ext}(y)$ such that³ $\operatorname{ext}(y)_j = y_i$ if j belongs to the i^{th} set in the collection I(and $\operatorname{ext}(y)_j = 0$ if $j \in [n] \setminus X_I$). Effectively, we query the function $F : \{0,1\}^n \to \{0,1\}$ defined as $F(x) = f(\operatorname{ext}(x_{\Gamma}))$, and this makes sense provided that F is close to f (under the distribution \mathcal{D}). To test the latter hypothesis condition, we sample \mathcal{D} and for each sample point x we compare f(x) to F(x), where here we again use the ability to determine the value of the influential variable in each set. Specifically, $\operatorname{ext}(x_{\Gamma})$ is computed by determining the value of x_{Γ} (without knowing Γ), and using our knowledge of $(X_i)_{i \in I}$.

We warn that the foregoing description presumes that we have correctly identified the collection I of all sets containing an influential variable. This leaves us with two questions: The first question is, how do we identify the set I. (Note that the influence of a variable may be as low as 2^{-k} , whereas we seek algorithms of poly(k)-complexity.) The solution (to be presented in Section 2.3.1) will be randomized, and will have one-sided error; specifically, we may fail to identify some sets that contain influential variables, but will never include in our collection sets that have no influential variables. Consequently, $f(1_{X_i} \circ 0_{\overline{X_i}}) \neq f(0^n)$ may not hold for some $i \in I$, and (over-simplifying again) we shall seek instead some $v^{(i)} \in \{0, 1\}^n$ such that $f(v^{(i)}) \neq f(w^{(i)})$, where $w^{(i)} = v_{\overline{X_i}}^{(i)} \circ 0_{X_i}$ (i.e., $w_j^{(i)} = v_j^{(i)}$ if $j \in [n] \setminus X_i$ and $w_j^{(i)} = 0$ otherwise). Second, as before, for every $i \in I$ and $x \in \{0, 1\}^n$, we wish to determine the value in x of the influential variable in the set X_i (a variable whose location is unknown to us!). This is done by observing that if $f \in C$ then this value equals $f(x') + f(v^{(i)}) + 1$ where $x' = x_{X_j} \circ v_{\overline{X_j}}^{(i)}$ (i.e., $x'_j = x_j$ if $j \in X_i$ and $x'_j = v_j^{(i)}$ otherwise).⁴ Again, we need to test whether $x \mapsto f(x') + f(v^{(i)}) + 1$ is a dictatorship, and use self-correction.

2.3 The Actual Tester

As warned, the above description is an over-simplification, and the actual way in which the set I is identified and used is more complex.

We fix a random partition of [n] to $r = O(k^2)$ parts, denoted (X_1, \ldots, X_r) . If $f \in C$, then, with high probability, each X_i contains at most one influential variable, denoted $\tau(i)$. We assume that this is the case when providing intuition throughout this section.

2.3.1 Stage 1: Finding I and corresponding $v^{(i)}$

Our goal is to find a collection I of at most k sets such that the function h_I is $\epsilon/3$ -close to f(w.r.t distribution \mathcal{D}), where h_I is defined as $h_I(x) = f(x_{X_I} \circ 0_{\overline{X_I}})$ and $X_I = \bigcup_{i \in I} X_i$. In addition, for each $i \in I$, we seek a witness $v^{(i)}$ for the fact that f depends on some variable in X_i ; that is, $f(v^{(i)}) \neq f(w^{(i)})$ for some $v^{(i)}$ that differ from $w^{(i)}$ only on X_i .

³ Notice that $\operatorname{ext}(y) = 0_{\overline{X_I}} \circ \left(\underset{i \in I}{\circ} (y_i)_{X_i} \right)$ - Here $(y)_X = 1_X$ if y = 1 and 0_X if y = 0.

⁴ Indeed, if $\tau(i) \in X_i$ is the index of the (unique) influential variable that resides in the set X_i , then

$$f(x') = x_{\tau(i)} \cdot f(v^{(i)}) + (x_{\tau(i)} + 1) \cdot f(w^{(i)}) = x_{\tau(i)} + f(v^{(i)}) + 1$$

since $f(v^{(i)}) + f(w^{(i)}) = 1$.

The procedure

We proceed in iterations, starting with $I = \emptyset$.

1. We sample \mathcal{D} for $O(1/\epsilon)$ times, trying to find $u \in \mathcal{D}$ such that $f(u) \neq h_I(u)$. (Note that if $I = \emptyset$, then $h_I(u) = f(0^n)$. In general, we seek u such that $f(u) \neq f(u_{X_I} \circ 0_{\overline{X_I}})$. If no such u is found, then we set $h = h_I$ and proceed to Stage 2. In this case, we may

If no such u is found, then we set $h = h_I$ and proceed to Stage 2. In this case, we may assume that h_I is $\epsilon/3$ -close to f (w.r.t \mathcal{D}).

2. Otherwise (i.e., $f(u) \neq h_I(u)$), we find an $i \in [m] \setminus I$ and $v^{(i)}$ such that $h_I(v^{(i)}) \neq h_{I \cup \{i\}}(v^{(i)})$, which means that X_i contains an influential variable and $v^{(i)}$ is the witness for the sensitivity that we seek. We set $I \leftarrow I \cup \{i\}$ and proceed to the next iteration. (We find this *i* by binary search that seeks *i* and *S* such that $h_{I \cup S \cup \{i\}}(u) \neq h_{I \cup S}(u)$, which means that $v^{(i)}$ equals *u* in locations outside *S* and is zero on *S*.)⁵

Once the iterations are suspended (due to not finding u), we reject if |I| > k, and continue to the Stage 2 otherwise. Recall that in the latter case $h = h_I$ is $\epsilon/3$ -close to f (w.r.t \mathcal{D}).

Note that if $f \in C$, then I contains only sets that contain variables of the k-junta, and so we never reject in this stage. In general, if $i \in I$, then $h_{I \setminus \{i\}}(v^{(i)}) \neq h_I(v^{(i)})$, which implies that $f(x') \neq f(x'')$, where x' and x'' differ only on X_i (e.g., $x''_{X_I} = v^{(i)}_{X_I}$ and $x''_j = 0$ if $j \notin X_I$).

2.3.2 Stage 2: Extracting the value of an influential variable

Given a collection I as found in Stage 1 (and a sensitivity witness $v^{(i)}$ for each $i \in I$), let $h = h_I$ and recall that h is close to f w.r.t \mathcal{D} . For each $i \in I$, given $x \in \{0, 1\}^n$, we wish to determine the value of x at the influential variable that resides in X_i .

For each $i \in I$, we define $\nu_i : \{0,1\}^{|X_i|} \to \{0,1\}$ such that $\nu_i(z) = h_I(y)$, where $y_{X_i} = z$ and $y_{\overline{X_i}} = v_{\overline{X_i}}^{(i)}$. Suppose that $f \in C$, and recall that $\tau(i) \in X_i$ denotes the location of the influential variable in X_i . Let $\sigma(i)$ denote the index of $\tau(i)$ in X_i (i.e., the $\sigma(i)^{\text{th}}$ element of X_i is $\tau(i)$). Then, in this case, ν_i is either a dictatorship or an anti-dictatorship. In particular, if ν_i is a dictatorship, then $\nu_i(z) = z_{\sigma(i)}$ (and otherwise $\nu_i(z) = z_{\sigma(i)} + 1$).

For each $i \in I$, we test whether ν_i is a dictatorship or anti-dictatorship, where testing is w.r.t the uniform distribution over $\{0,1\}^{|X_i|}$. Note that we also check whether ν_i is a dictatorship or anti-dictatorship. If the tester (run with proximity parameter 0.1) fails, we reject. Otherwise (i.e., if we did not reject), we can compute ν_i via self-correction on h_I ; that is, to compute ν_i at z, we select $u \in \{0,1\}^{|X_i|}$ at random, and return $\nu_i(z+u) + \nu_i(u)$, which (w.h.p.) equals $(z+u)_{\sigma(i)} + u_{\sigma(i)} = z_{\sigma(i)}$.

Hence, we always continue to Stage 3 if $f \in C$, and whenever we continue to Stage 3 we can compute all ν_i (for $i \in I$) via self-correction.

2.3.3 Stage 3: Emulating a tester of C'

Recall that when reaching this stage, we may assume that $h = h_I$ is $\epsilon/3$ -close to f (w.r.t \mathcal{D}). Also recall that $h_I(x)$ depends only on x_{X_I} , where $X_I = \bigcup_{i \in I} X_i$, and that by Stage 2 we may assume that $\nu_i(z) = z_{\sigma(i)}$ (for every $i \in I$ and almost all z). In light of the forgoing, we define $F : \{0,1\}^n \to \{0,1\}$ such that F(x) = h(x') where $x'_{X_i} = (x_{\sigma(i)}, \ldots, x_{\sigma(i)})$ (i.e.,

⁵ By Step 1, we have $h_{S'\cup I}(u) \neq h_{S''\cup I}(u)$, for $S' = [n] \setminus I$ and $S'' = \emptyset$, and in each iteration we cut $S' \setminus S''$ by half while maintaining $h_{S'\cup I}(u) \neq h_{S''\cup I}(u)$.

 $x'_j = (x_{X_i})_{\sigma(i)} = x_{\tau(i)}$ if $j \in X_i)^6$ and $x'_j = 0$ otherwise. (Indeed, if $f \in C$, then F(x) = h(x), since h(y) depends only on $(y_{\tau(i)})_{i \in I}$. Using hypothesis that C' (and so C) is closed under zero projection, it follows that $F \in C$.)

We observe that if F is $\epsilon/3$ -close (w.r.t \mathcal{D}) to both h and C, then f must be ϵ -close to C (since f is $\epsilon/3$ -close to h). Hence, we test both these conditions. Specifically, using our ability to sample \mathcal{D} , query f, and determine the value of the influential variables in X_I , we proceed as follows:

1. Test whether F = h, where testing is w.r.t the distribution \mathcal{D} and proximity parameter $\epsilon/3$.

This is done by taking $O(1/\epsilon)$ samples of \mathcal{D} , and comparing the values of F and h on these sample points. Recall that $h(u) = h_I(u) = f(u_{X_I} \circ 0_{\overline{X_I}})$.

The value of F on u is determined as follows.

- a. For every $i \in I$, if ν_i is a dictatorship, then set v_i to equal the self-corrected value of $\nu_i(u_{X_i})$, where ν_i is as defined in Stage 2. Otherwise (i.e., when ν_i is an anti-dictatorship), we set v_i to equal the self-corrected value of $\nu_i(u_{X_i}) + 1$.
- **b.** Return the value h(u'), where $u'_j = v_i$ if $j \in X_i$ and $u'_j = 0$ otherwise.

Indeed, F = h always passes this test, whereas F that is $\epsilon/3$ -far from h (w.r.t \mathcal{D}) is rejected w.h.p.

2. Test whether F is in C, where testing is w.r.t the distribution \mathcal{D} and proximity parameter $\epsilon/3$. This is done by testing whether F' is in C, where F'(z) = F(x) such that $x_j = z_i$ if j is in the i^{th} set in the collection I, and $x_j = 0$ otherwise. Here we use a distribution-free tester, and analyze it w.r.t the distribution \mathcal{D}_I . Toward this end, we need to samples \mathcal{D}_I as well as answer queries to F', where both tasks can be performed as in the prior step. Recall that if $f \in C$, then $F \in C$, and this test will accept (w.h.p.), whereas if F is $\epsilon/3$ -far from C the test will reject (w.h.p.).

We conclude that if we reached Stage 3 and $f \in C$ (resp., f is ϵ -far from C), then we accept (resp., reject) w.h.p.

2.4 Digest: Our approach vs the original one [24]

Our new approach differs from the original approach of Diakonikolas et al. [24] in two main aspects:

- 1. In [24], sets that contain influential variables are identified according to their influence, which is defined with respect to the uniform distribution. This definition seems inadequate when dealing with arbitrary distributions. Instead, we identify such a set by searching for two assignments that differ only on this set and yield different function values. The actual process is iterative and places additional constraints on these assignments (as detailed in Section 2.3.1).
- 2. In [24], given an assignment to the function, the value of the unique influential variable that resides in a given set S is determined by approximating the influence of two subsets of S (i.e., the subsets of locations assigned the value 0 and 1, respectively). In contrast, we determine this value by defining an auxiliary function, which depends on the unknown influential variable, and evaluating this function (via self-correction w.r.t the uniform distribution (!); see Section 2.3.2).

⁶ In general, $\tau(i)$ denotes the location in [n] of the $\sigma(i)^{\text{th}}$ element of X_i .

2.5 The Model

In this subsection, we define the testing and learning models.

In the testing model, we consider the problem of testing a Boolean function class Cin the uniform and distribution-free testing models. In the distribution-free testing model (resp. uniform model), the algorithm has access to a Boolean function f via a black-box that returns f(x) when a string x is queried. We call this query membership query (MQ_f or just MQ). The algorithm also has access to unknown distribution \mathcal{D} (resp. uniform distribution) via an oracle that returns $x \in \{0, 1\}^n$ chosen randomly according to the distribution \mathcal{D} (resp. according to the uniform distribution). We call this query example query (ExQ_{\mathcal{D}} (resp. ExQ or ExQ_U)).

A distribution-free testing algorithm, [32], (resp. testing algorithm) \mathcal{A} for C is an algorithm that, given as input a distance parameter ϵ and the above two oracles to a Boolean function f, 1. if $f \in C$ then \mathcal{A} outputs "accept" with probability at least 2/3.

2. if f is ϵ -far from every $g \in C$ with respect to the distribution \mathcal{D} (resp. uniform distribution) then \mathcal{A} outputs "reject" with probability at least 2/3.

We will also call \mathcal{A} a tester (or ϵ -tester) for the class C and an algorithm for ϵ -testing C. We say that \mathcal{A} is one-sided if it always accepts when $f \in C$; otherwise, it is called two-sided algorithm. The query complexity of \mathcal{A} is the maximum number of queries \mathcal{A} makes on any Boolean function f.

In the learning models, C is a class of representations of Boolean functions rather than a class of Boolean functions. Therefore, we may have two different representations in Cthat are logically equivalent. In this paper, we assume that this representation is verifiable; that is, given a representation g, one can decide in polynomial time on the length of this representation if $g \in C$.

A distribution-free proper learning algorithm (resp. proper learning algorithm under the uniform distribution) \mathcal{A} for C is an algorithm that, given as input an accuracy parameter ϵ , a confidence parameter δ and an access to both MQ_f for the target function $f \in C$ and ExQ_D, with unknown \mathcal{D} , (resp. ExQ or ExQ_U), with probability at least $1 - \delta$, \mathcal{A} returns $h \in C$ that is ϵ -close to f with respect to \mathcal{D} (resp. with respect to the uniform distribution). This model is also called proper PAC-learning with membership queries under any distribution (resp. under the uniform distribution) [3, 48].

3 The Distribution-Free Tester

In this section, we sketch the proof of the tester from Section 2.

For a class C of n-bit long Boolean functions and a set $Y = \{y_1, \ldots, y_q\}$ we define $C^*(Y)$ the class of all q-bit long Boolean functions $f(y_1, \ldots, y_q) = g(y_1, \ldots, y_q, 0, \ldots, 0)$ where $g \in C$. We define $C(Y) \subseteq C^*(Y)$ the class of $f \in C(Y)$ that depends on all the variables in Y. That is, all the variables in Y are influential.

Our main result is

▶ **Theorem 1.** Let C be a class of n-bit long Boolean functions that is symmetric subclass of k-Junta and is closed under zero projection. Suppose for every $Y = \{y_1, \ldots, y_q\}$ with $q \leq k$, there is a tester T_Y for q-bit Boolean function F such that

- 1. T_Y is a polynomial time two-sided distribution-free (resp. uniform-distribution) adaptive ϵ -tester
- **2.** If $F \in C(Y)$ then, with probability at least 1δ , T_Y accepts.

- **3.** If F is ϵ -far from every function in $C^*(Y)$ w.r.t \mathcal{D} then, with probability at least 1δ , T_Y rejects.
- **4.** T_Y makes $M(\epsilon, \delta)$ MQs and $Q(\epsilon, \delta)$ $ExQ_{\mathcal{D}}$ (resp. ExQ_U).

Then

1. There is a polynomial time two-sided distribution-free adaptive algorithm for ϵ -testing C that makes

$$\tilde{O}\left(M(\epsilon/12, 1/24) + kQ(\epsilon/12, 1/24) + \frac{k}{\epsilon}\right)$$

queries.

2. (resp. There is a polynomial time two-sided uniform-distribution adaptive algorithm for ϵ -testing C that makes

$$\tilde{O}\left(M(\epsilon/12, 1/24) + Q(\epsilon/12, 1/24) + \frac{k}{\epsilon}\right)$$

queries.)

Using Goldreich et. al [32], reduction of testing to proper learning we get

▶ **Theorem 2.** Let C be a class of n-bit long Boolean functions that is symmetric subclass of k-Junta and is closed under zero projection. Suppose for every $Y = \{y_1, \ldots, y_q\}$ with $q \leq k$, there is a polynomial time proper learning algorithm that learns $C^*(Y)$ using $M(\epsilon, \delta)$ MQs and $Q(\epsilon, \delta)$ ExQ_Ds (resp. ExQ_Us). Then

1. There is a polynomial time two-sided distribution-free adaptive algorithm for ϵ -testing C that makes

$$\tilde{O}\left(M(\epsilon/12, 1/24) + kQ(\epsilon/12, 1/24) + \frac{k}{\epsilon}\right)$$

queries.

2. (resp. There is a polynomial time two-sided uniform-distribution adaptive algorithm for ϵ -testing C that makes

$$\tilde{O}\left(M(\epsilon/12, 1/24) + Q(\epsilon/12, 1/24) + \frac{k}{\epsilon}\right)$$

queries.)

Before we give the proof sketch, we give some applications.

3.1 Some Applications

- *k*-Junta: For the class C = k-Junta, $C^*(Y) = q$ -Junta and since f is q-bit Boolean function $f \in C^*(Y)$ and the tester T_Y can just returns accept. Then M = Q = 0 and we get a distribution-free tester for k-Junta that makes $\tilde{O}(k/\epsilon)$.
- *k*-Linear: For the class C = k-Linear, the sum (over F_2) of at most k variables, we have $C(Y) = \{y_1 + y_2 + \dots + y_q\}$. If f is ϵ -far from $C^*(Y)$ then it is ϵ -far from $y_1 + y_2 + \dots + y_q$. We can distinguish between $g = y_1 + y_2 + \dots + y_q$ and a function that is ϵ -far from g with $O(1/\epsilon) \operatorname{ExQ}_{\mathcal{D}}$. Then M = 0 and $Q = O(1/\epsilon)$. So we get a distribution-free tester for k-Linear that makes $O(k/\epsilon)$ queries.

- *k*-Term: For the class C = k-Term, the conjunction of at most k literals, we have $C(Y) = \{z_1 \land z_2 \land \cdots \land z_q \mid z_i \in \{y_i, \overline{y_i}\}\}$. The tester T_Y asks $O(1/\epsilon) \operatorname{ExQ}_{\mathcal{D}}$. If for all the strings f is zero then the tester accept. Otherwise, there is a string a such that f(a) = 1. Then $z_i = y_i$ if $a_i = 1$ and $z_i = \overline{y_i}$ if $a_i = 0$. Then as above, it tests if f is ϵ -far from $z_1 \land z_2 \land \cdots \land z_q$ and we get a distribution-free tester for k-Term that makes $O(k/\epsilon)$ queries.
- s-term monotone r-DNF: For the class s-term monotone r-DNF (a DNF that contains at most s r-Terms), in the full paper, we give an algorithm that properly learns this class in polynomial time and makes $O(s/\epsilon) \operatorname{ExD}_{\mathcal{D}}$ and $O(rs \log(ns))$ MQs. Since the number of influential variables in any s-term monotone r-DNF is at most sr we have $q \leq k = sr$. Therefore the class $C^*(Y)$ $(n = q \leq sr)$ can be properly learned using $O(s/\epsilon) \operatorname{ExD}_{\mathcal{D}}$ and $O(rs \log(rs))$ MQs. By Theorem 2 we get

▶ **Theorem 3.** For any $\epsilon > 0$, there is a polynomial time two-sided distribution-free adaptive algorithm for ϵ -testing s-Term Monotone r-DNF that makes $\tilde{O}(rs^2/\epsilon)$ queries. In the full paper, we also show that

▶ **Theorem 4.** For any $\epsilon > 0$, there is a polynomial time two-sided distribution-free adaptive algorithm for ϵ -testing s-Term Unate⁷ r-DNF that makes $\tilde{O}(rs^2/\epsilon)$ queries.

3.2 Proof Sketch

Algorithm 1 A distribution-free tester for subclasses *C* of *k*-Junta.

Tester $C(f, \mathcal{D}, \epsilon)$ Input: Oracle that accesses a Boolean function f and \mathcal{D} . Output: If any one of the procedures rejects then "reject", otherwise, "accept"

- 1. $(X, V, I) \leftarrow \operatorname{ApproxTarget}(f, \mathcal{D}, \epsilon, 1/3).$
- 2. Define $h(x) = f(x_X \circ 0_{\overline{X}})$.
- 3. $\mathbf{TestSets}(h, X, V, I).$
- 4. Define F
- 5. $\mathbf{Close} fF(f, \mathcal{D}, \epsilon, 1/15)$
- 6. Run T_Y on F
- 7. Accept

We now sketch the proof of Theorem 1. See the full proof in [13]. The tester is **Tester**C in Algorithm 1. First, **Tester**C calls the procedure **ApproxTarget**, in Algorithm 2. This procedure executes the first stage of the tester. See Subsection 2.3.1. The reason that here the procedure is more complex is because, for learning classes of unate functions the procedure in Algorithm 1 returns witnesses for $h(x) = f(x_{X_I} \circ 0_{\overline{X_I}})$ and not for f (as in Subsection 2.3.1). So those witnesses also give us the unateness of each variable in the function. This is, for example, how we get the result in Theorem 4.

Throughout the paper we denote $X = X_I$ and $h(x) = f(x_X \circ 0_{\overline{X}})$. In [13] we prove the following. The proof sketch is in Subsection 2.3.1.

⁷ The function f is unate if there is a such that f(x + a) is monotone.

Algorithm 2 A procedure that finds influential sets $\{X_i\}_{i \in I}$ of f and a witness $v^{(i)}$ for each influential set X_i for $h := f(x_{X_I} \circ 0_{\overline{X_I}})$ where $X_I = \bigcup_{i \in I} X_i$. Also, whp, h is $(\epsilon/3)$ -close to the target.

ApproxTarget $(f, \mathcal{D}, \epsilon, c)$ Input: Oracle that accesses a Boolean function f and an oracle that draws $x \in \{0, 1\}^n$ according to the distribution \mathcal{D} . Output: Either "reject" or (X_I, V, I)

Partition [n] into r sets

1. Set $r = 2k^2$.

2. Choose uniformly at random a partition X_1, X_2, \ldots, X_r of [n]

Find a close function and influential sets

3. Set $X_I = \emptyset$; $I = \emptyset$; $V = \emptyset$; $t(X_I) = 0$. 4. Repeat $M = 3k \ln(15k)/\epsilon$ times 5.Choose $u \in \mathcal{D}$. 6. $t(X_I) \leftarrow t(X_I) + 1$ If $f(u_{X_I} \circ 0_{\overline{X_I}}) \neq f(u)$ then 7. $W \leftarrow \emptyset.$ 8. $(\ell, w^{(\ell)}) \leftarrow$ Binary Search to find a new influential set X_{ℓ} 9. using u and $u_{X_I} \circ 0_{\overline{X_I}}$ and a string $w^{(\ell)} \in \{0,1\}^n$ such that $f(w^{(\ell)}) \neq f(w_{\overline{X_{\ell}}}^{(\ell)} \circ 0_{X_{\ell}});$ 10. $X_I \leftarrow X_I \cup X_\ell; I \leftarrow I \cup \{\ell\}.$ 11. If |I| > k then Output("reject"). 12. $W = W \cup \{w^{(\ell)}\}.$ 13.Choose $w^{(r)} \in W$. 14.If $f(w_{X_I}^{(r)} \circ 0_{\overline{X_I}}) \neq f(w_{X_I \setminus X_r}^{(r)} \circ 0_{\overline{X_I} \cup X_r})$ then $W \leftarrow W \setminus \{w^{(r)}\}; v^{(r)} \leftarrow w_{X_I}^{(r)} \circ 0_{\overline{X_I}}; V \leftarrow V \cup \{v^{(r)}\};$ 15.If $W \neq \emptyset$ then Goto 14 Else If $f(w_{X_I}^{(r)} \circ 0_{\overline{X_I}}) \neq f(w^{(r)})$ then $u \leftarrow w^{(r)}$; Goto 9 Else $u \leftarrow w_{\overline{X_r}}^{(r)} \circ 0_{X_r}$; Goto 9 16.17.18. $t(X_I) = 0.$ 19. If $t(X_I) = c \ln(15k)/\epsilon$ then $\text{Output}(X_I, V, I)$.

▶ Lemma 5. Consider steps 1-2 in the Approx Target. If f is a k-junta then, with probability at least 2/3, for each $i \in [r]$, the set $x(X_i) = \{x_j | j \in X_i\}$ contains at most one influential variable of f.

Lemma 6. If Approx Target does not reject then it outputs (X, V, I) that satisfies

- 1. $q = |I| \le k$.
- **2.** For every $\ell \in I$, $v_{\overline{X}}^{(\ell)} = 0$ and $f(v^{(\ell)}) \neq f(0_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$. That is, $v^{(\ell)} \in V$ is a witness of $h(x) = f(x_X \circ 0_{\overline{X}})$ for X_{ℓ} .
- **3.** Each $x(X_{\ell}), \ell \in I$, contains at least one influential variable of $h(x) = f(x_X \circ 0_{\overline{X}})$.
- 4. With probability at least 14/15

$$Pr_{u\in\mathcal{D}}[h(x)\neq f(x)]\leq\epsilon/3.$$

▶ Lemma 7. If f is k-junta and each $x(X_i)$ contains at most one influential variable of f then

- 1. Approx Target outputs (X, V, I).
- **2.** Each $x(X_{\ell}), \ell \in I$, contains exactly one influential variable in $h(x) = f(x_X \circ 0_{\overline{X}})$.
- **3.** For every $\ell \in I$, $h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)}) = f(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is a literal.

▶ Lemma 8. The procedure ApproxTarget makes $O((k \log k)/\epsilon) = \tilde{O}(k/\epsilon)$ queries.

The tester then defines $h(x) = f(x_X \circ 0_{\overline{X}})$. Now, since $\mathbf{Pr}_{u \in \mathcal{D}}[h(x) \neq f(x)] \leq \epsilon/3$ it is enough to distinguish between h is in C and h that is $(2\epsilon/3)$ -far from every function in Cwith respect to \mathcal{D} .

The tester then moves to the second stage. See Subsection 2.3.2. First, it calls the procedure **TestSets**. See Algorithm 3 in Appendix A. The procedure tests if every $h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is close to a literal. In the procedure, **UniformJunta** (g, k, ϵ, δ) is Blais's uniform-distribution one-sided tester [7] for k-Junta. For k = 1 it tests whether g is a literal or constant function or ϵ -far from any literal and constant function with respect to the uniform distribution.

The following is very easy to prove

- ► Lemma 9. We have
- 1. If h is k-junta and each $x(X_i)$ contains at most one influential variable of f then **TestSets** returns "OK".
- 2. If for some $\ell \in I$, $h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is (1/30)-far from every literal with respect to the uniform distribution then, with probability at least 1 (1/15), **TestSets** rejects.
- **3.** The procedure **TestSets** makes O(k) queries.

This test does not give $\tau(i)$ (the index of the influential variable in X_i) but the fact that $h(x_{X_i} \circ v_{\overline{X_i}}^{(i)})$ is close to $x_{\tau(i)}$ or $\overline{x_{\tau(i)}}$ can be used to find the value of $u_{\tau(i)}$ in every assignment $u \in \{0,1\}^n$ without knowing $\tau(i)$. The latter is done, whp, by the procedure **RelVarValues** that uses self-correction. See Algorithm 4 in Appendix A. We have

- ► Lemma 10. We have
- 1. If h is k-Junta and each $x(X_i)$ contains at most one influential variable of h then **RelVarValues** outputs z such that $z_{\ell} = w_{\tau(\ell)}$ where $h(x_{X_{\ell}} \circ 0_{\overline{X_{\ell}}}) \in \{x_{\tau(\ell)}, \overline{x_{\tau(\ell)}}\}.$
- **2.** If for every $\ell \in I$ the function $h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is (1/30)-close to a literal in $\{x_{\tau(\ell)}, \bar{x}_{\tau(\ell)}\}$ with respect to the uniform distribution, where $\tau(\ell) \in X_{\ell}$, and **RelVarValues** does not reject then, with probability at least 1δ , we have: For every $\ell \in I$, $z_{\ell} = w_{\tau(\ell)}$.
- **3.** The procedure **RelVarValues** makes $O(k \log(k/\delta))$ queries.

Proof. Since $\nu_i(x) = h(x_{X_\ell} \circ v_{\overline{X_\ell}}^{(\ell)})$ is (1/30)-close to a literal in $\{x_{\tau(\ell)}, \overline{x}_{\tau(\ell)}\}$ we have that for a uniform random string z, with probability at least 1/15 we have $\nu(x+z) + \nu(z) = x_{\tau(\ell)}$. If we repeat this test $O(\log(k/\delta))$ times for every $\ell \in I$, we get a success probability $1 - \delta$. Since $|I| \leq k$, **RelVarValues** makes $O(k \log(k/\delta))$ queries.

We collect all the above events that happens with high probability in the following

▶ Assumption 11. For the rest of this section we assume

1. If $f \in C$ then

- $h(x) = f(x_X \circ 0_{\overline{X}}) \in C.$
- Each $x(X_{\ell}), \ell \in I$ contains exactly one influential variable.
- For every $\ell \in I$, $f(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is a literal in $\{x_{\tau(\ell)}, \overline{x_{\tau(\ell)}}\}$.

- **2.** If f is ϵ -far from every function in C then
 - $h(x) = f(x_X \circ 0_{\overline{X}})$ is $(\epsilon/3)$ -close to f with respect to \mathcal{D} and therefore h(x) is $(2\epsilon/3)$ -far from every function in C with respect to \mathcal{D} .
 - For every $\ell \in I$, $f(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is (1/30)-close to a literal in $\{x_{\tau(\ell)}, \bar{x}_{\tau(\ell)}\}$ with respect to the uniform distribution.
- **3.** For any u, we can get $(u_{\tau(\ell)})_{\ell \in I}$ using $\tilde{O}(k)$ queries.

We are getting now to the third stage. See Subsection 2.3.3. For a function $f \in C$ we define $\operatorname{Rel}(f)$, the set of all influential variables of f. Let q = |I| and $\Gamma := \{\tau(\ell_1), \ldots, \tau(\ell_q)\}$. Notice that, with the above assumption, if $h \in C$ then $\operatorname{Rel}(h) = \Gamma$. We define the class $C(\Gamma)$ (resp. $C^*(\Gamma)$), the set of functions in C with $\operatorname{Rel}(f) = \{x_\gamma \mid \gamma \in \Gamma\}$ (resp. $\operatorname{Rel}(f) \subseteq \{x_\gamma \mid \gamma \in \Gamma\}$). Since C is symmetric and closed under zero projection $C(\Gamma)$ (resp. $C^*(\Gamma)$) is the set of all functions $f(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)}, 0, 0, \ldots, 0)$ where $f \in C$ and $\operatorname{Rel}(f) = \{x_1, \ldots, x_q\}$ (resp. $\operatorname{Rel}(f) \subseteq \{x_1, \ldots, x_q\}$). We recall that, for $Y = \{y_1, \ldots, y_q\}$, C(Y) (resp. $C^*(Y)$) are the set of all functions $f(y_1, \ldots, y_q, 0, 0, \ldots, 0)$ where $f \in C$ and $\operatorname{Rel}(f) = \{x_1, \ldots, x_q\}$ (resp. $\operatorname{Rel}(f) \subseteq \{x_1, \ldots, x_q\}$).

Let $F(y_1, \ldots, y_q) = h(z)(=f(z))$ where $z = (y_1)_{X_{\ell_1}} \circ \cdots \circ (y_q)_{X_{\ell_q}} \circ 0_{\overline{X}}$. That is, for every $\ell_i \in I$, $j \in X_{\ell_i}$ we have $z_j = y_i$ and for every $j \in \overline{X}$ we have $z_j = 0$. Then, by Assumption 11, it is easy to prove that (see Subsection 2.3.3)

► Lemma 12. We have

- **1.** If $h \in C$ then $F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)}) = h$.
- **2.** If $h \in C$ then $F(y_1, \ldots, y_q) \in C(Y)$.
- 3. If h is 2ε/3-far from every function in C with respect to D then either
 a. h is ε/3-far from F(x_{τ(ℓ1}),...,x_{τ(ℓq})) with respect to D, or,
 - **b.** $F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)})$ is $\epsilon/3$ -far from every function in C with respect to \mathcal{D} .

Therefore, we need to do two tests. The first is to distinguish between h = F and h that is $\epsilon/3$ -far from F w.r.t \mathcal{D} . The second is to distinguish between $F \in C$ and F that is $\epsilon/3$ -far from every function in C w.r.t \mathcal{D} .

The following result shows that we can query $F(y_1, \ldots, y_q)$ and for every $x \in \{0, 1\}^n$ we can extract x_{Γ} . So, in particular, we can get a sample according to \mathcal{D}_{Γ} and query $F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)})$. The proof is immediate (See Appendix A)

Lemma 13. For the function F we have

- 1. Given (y_1, \ldots, y_q) , computing $F(y_1, \ldots, y_q)$ can be done with one query to f.
- **2.** Given $x \in \{0,1\}^n$ and δ , there is an algorithm that makes $O(k \log(k/\delta))$ queries and, with probability at least 1δ , either discovers that $f \notin C$ and then reject or computes $x_{\Gamma} = (x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)})$ and $F(x_{\Gamma})$.
- 3. There is a polynomial time algorithm that makes $O(k \log(k/\delta))$ queries and with probability at least $1 - \delta$ returns a string u in $\{0, 1\}^q$ according to the distribution \mathcal{D}_{Γ} and F(u).
- 4. There is a polynomial time algorithm that makes one query and returns a string u in $\{0,1\}^q$ according to the uniform distribution and F(u).

We now give the procedure $\operatorname{Close} fF$ that tests whether $h(x) = f(x_X \circ 0_{\overline{X}})$ is $(\epsilon/3)$ -far from F with respect to \mathcal{D} . See Algorithm 5 and the proof in Appendix A.

▶ Lemma 14. For any ϵ , a constant δ , and (X, V, I) that satisfies Assumption 11, procedure Closef F makes $O((k/\epsilon)\log(k/\epsilon)) = \tilde{O}(k/\epsilon)$ queries and

- 1. If $f \in C$ then ClosefF returns OK.
- 2. If h(x) is $(\epsilon/3)$ -far from F with respect to D then, with probability at least $1-\delta$, Closef F rejects.

The second test is to distinguish between $F \in C$ and F that is $\epsilon/3$ -far from C w.r.t the distribution \mathcal{D} .

Consider the tester T_Y in Theorem 1. If $h(x) = F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)}) \in C$ then, by Assumption 11, since each $x(X_\ell), \ell \in I$, contains exactly one influential variable, $F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)}) \in C(\Gamma)$. Therefore $F = F(y_1, \ldots, y_q) \in C(Y)$. If $F(x_{\tau(\ell_1)}, \ldots, x_{\tau(\ell_q)})$ is $\epsilon/3$ -far from every function in C w.r.t \mathcal{D} , then it is $\epsilon/3$ -far from every function in $C^*(\Gamma)$ w.r.t \mathcal{D}_{Γ} . Therefore, $F = F(y_1, \ldots, y_q)$ is $\epsilon/3$ -far from every function in $C^*(\Gamma)$ w.r.t \mathcal{D}_{Γ} . Therefore, $F = F(y_1, \ldots, y_q)$ is $\epsilon/3$ -far from every function in $C^*(Y)$ w.r.t \mathcal{D}_{Γ} . Therefore, T_Y can be used for the second test of $F(x_{\Gamma}) \in C$ vs. $\epsilon/3$ far from every function in C.

Now by, Lemma 13, every MQ to F can be simulated with one MQ to f and every $\operatorname{ExQ}_{\mathcal{D}_{\Gamma}}$ can be simulated with $O(k \log(k/\delta')) \operatorname{ExQ}_{\mathcal{D}}$ queries and have success probability $1-\delta'$. Since T_Y asks $Q(\epsilon, \delta)$ queries, we need $O(k \log(kQ(\epsilon, \delta)/\delta)) \operatorname{ExQ}_{\mathcal{D}}$ to have success $1-\delta$ for all the queries. When \mathcal{D} is uniform then, by Lemma 13, every $\operatorname{ExQ}_{U_{\Gamma}}$ can be simulated with one query to MQ.

Notice that the success probability in all the procedures above is $1 - \delta$ for any constant δ . By choosing $\delta = 1/20$ for each procedure, we get a tester with confidence of at least 2/3. By Lemmas 8, 9 and the above analysis the query complexity of the tester is as stated in Theorem 1.

4 The Second Tester: Uniform-Distribution Tester for Classes that are Close to *k*-Junta w.r.t the Uniform Distribution

The second tester in this paper tests classes that are close to k-Junta w.r.t the uniform distribution, that is, for every $f \in C$ and every ϵ , there is k such that every function in C is ϵ -close to some function in k-Junta.

To understand the intuition behind the second tester, we demonstrate it for testing s-term DNF, the class of DNF with at most s terms. This class is close to $(s \log(s/\epsilon))$ -Junta. This is because, for every s-term DNF and every ϵ , the function g that results by removing the terms of size greater than or equal to $\log(s/\epsilon)$ in f is ϵ -close to f and $g \in (s \log(s/\epsilon))$ -Junta.

Let f be an s-term DNF. Since, for ϵ -testing, the variables that are influential in f are variables in terms of size $d = O(\log(s/\epsilon))$, there are at most $k = sd = O(s\log(s/\epsilon))$ influential variables in f. Suppose we uniformly at random distribute the variables of f into r = 10k bins X_1, \ldots, X_r . The influential variables falls into at most k bins. We call those bins influential bins. Terms in f of size greater than d, with high probability, d/2 of their variables falls into the uninfluential bins.

We try as before to find the influential bins, but this time, we use uniform random strings in our search. This is because, when we use uniform random strings, with high probability, all the large terms in f are equal zero for those strings, and therefore, no uninfluential bin is found by the search procedure.

We find enough influential bins such that if we substitute a random assignment in the variables of the uninfluential bins, w.h.p., we get a function H that is $\epsilon/4$ -close to f. The next key idea is: as we said before, if we have a large term in f, then with high probability, many of its variables fall into the uninfluential bins. So when we substitute a random assignment for the variables in the uninfluential bins, with high probability, the large terms in f vanish in H. Therefore, with high probability, H is $\epsilon/4$ close to f and contains small terms (terms of size at most $d = O(\log(s/\epsilon))$. Since H is s-term d-DNF, it is a function in sd-Junta, and we can use the first tester to test H.

See more details in Appendix B.

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5:16 Optimal Testers

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A The Procedures of the First Tester

Proof of Lemma 13.

Proof. 1 is immediate since $F(y_1, \ldots, y_q) = f(z)$ where $z = (y_1)_{X_{\ell_1}} \circ \cdots \circ (y_q)_{X_{\ell_q}} \circ 0_{\overline{X}}$. To prove 2. We run **RelVarValues** (x, X, V, I, δ) . If it rejects then, by Lemma 10, $f \notin C$. Otherwise, by Lemma 10, with probability at least $1 - \delta$, the procedure outputs z where for every ℓ , $z_{\ell} = x_{\tau(\ell)}$. Then using 1 we compute F(z). Since by Lemma 10, **RelVarValue** makes $O(k \log(k/\delta))$ queries, the result follows. Now 3 and 4 follows immediately from 1 and 2. **Algorithm 3** A procedure that tests if for all $\ell \in I$, $h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ is (1/30)-close to some literal with respect to the uniform distribution.

TestSets(h, X, V, I)Input: Oracle that accesses a Boolean function f and (X, V, I). Output: Either "reject" or "OK"

```
1. For every \ell \in I do

2. If UniformJunta(h(x_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)}), 1, 1/30, 1/15) = "reject"

3. then Output("reject")

4. Choose b \in \{0, 1\}^n uniformly at random

5. If h(b_{X_{\ell}} \circ v_{\overline{X_{\ell}}}^{(\ell)}) = h(\overline{b_{X_{\ell}}} \circ v_{\overline{X_{\ell}}}^{(\ell)}) then Output("reject")

6. Return "OK"
```

Algorithm 4 A procedure that takes as input (X, V, I) and a string $w \in \{0, 1\}^n$ and, with probability at least $1 - \delta$, returns the values of $w_{\tau(i)}$, $i \in I$, where $h(x_{X_i} \circ v_{\overline{X_i}}^{(i)})$ is (1/30)-close to one of the literals in $\{x_{\tau(i)}, \overline{x_{\tau(i)}}\}$ with respect to the uniform distribution.

$\mathbf{RelVarValues}(w, X, V, I, \delta)$

Input: Oracle that accesses a Boolean function h, (X, V, I) and $w \in \{0, 1\}^n$. *Output*: Either "reject" or returns for every $\ell \in I$, the value, $z_{\ell} = w_{\tau(\ell)}$ where $x_{\tau(\ell)}$ is one of the influential variables of $h(x_X \circ 0_{\overline{X}})$ in $x(X_{\ell})$

1.	For every $\ell \in I$ do
2.	For $\xi \in \{0, 1\}$ set $Y_{\ell, \xi} = \{j \in X_{\ell} w_j = \xi\}.$
3.	Set $G_{\ell,0} = G_{\ell,1} = 0;$
4.	Repeat $t = \ln(k/\delta)/\ln(4/3)$ times
5.	Choose $b \in \{0, 1\}^n$ uniformly at random;
6.	If $h(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v_{\overline{X_{\ell}}}^{(\ell)}) \neq h(\overline{b_{Y_{\ell,0}}} \circ b_{Y_{\ell,1}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ then $G_{\ell,0} \leftarrow G_{\ell,0} + 1$
7.	If $h(b_{Y_{\ell,1}} \circ b_{Y_{\ell,0}} \circ v_{\overline{X_{\ell}}}^{(\ell)}) \neq h(\overline{b_{Y_{\ell,1}}} \circ b_{Y_{\ell,0}} \circ v_{\overline{X_{\ell}}}^{(\ell)})$ then $G_{\ell,1} \leftarrow G_{\ell,1} + 1$
8.	If $(\{G_{\ell,0}, G_{\ell,1}\} \neq \{0, h\})$ then Output("reject")
9.	If $G_{\ell,0} = t$ then $z_{\ell} \leftarrow 0$ else $z_{\ell} \leftarrow 1$
10.	$\operatorname{Output}(``\{z_\ell\}_{\ell \in I}")$

Algorithm 5 A procedure that tests whether h(x) is $(\epsilon/3)$ -far from F with respect to \mathcal{D} .

Close $fF(f, \mathcal{D}, \epsilon, \delta)$ Input: Oracle that accesses a Boolean function f and \mathcal{D} . Output: Either "reject" or "OK"

- 1. Repeat $t = (3/\epsilon) \ln(2/\delta)$ times
- 2. Choose $u \in \mathcal{D}$.
- 3. $z \leftarrow \mathbf{RelVarValue}(u, X, V, I, \delta/(2t))$.
- 4. If $h(u) \neq F(z)$ then Output("reject")
- 5. Return "OK".

Proof of Lemma 14.

Proof. If $f \in C$ then, by Lemma 12, $F(u_{\Gamma}^{(i)}) = h(u)$ for every *i*. By Lemma 10 and Assumption 11, $z^{(i)} = u_{\Gamma}^{(i)}$ for all *i*, and therefore **Close** *f F* returns OK.

Suppose now h(x) is $(\epsilon/3)$ -far from F with respect to \mathcal{D} . By 2 in Lemma 13, **RelVarValue** makes $O(kt \log((kt)/\delta))$ queries and computes $F(u_{\Gamma}^{(i)})$, $i = 1, \ldots, t$, with failure probability at most $\delta/2$. Then the probability that it fails to reject is at most $(1 - \epsilon/3)^t \leq \delta/2$. This gives the result.

Therefore, **Close** fF makes $O((k/\epsilon)\log(k/\epsilon))$ queries and satisfies 1 and 2.

B Classes that are Close to k-Junta

We now give more details. The tester first runs the procedure **Approx**C in Algorithm 6. This procedure is similar to the procedure **ApproxTarget**. It randomly uniformly partitions the variables to $r = 4c^2(c+1)s \log(s/\epsilon)$ disjoint sets X_1, \ldots, X_r and finds influential sets $\{X_i\}_{i \in I}$. Here c is a large constant. To find a new influential set, it chooses two random uniform strings $u, v \in \{0, 1\}^n$ and verifies if $f(u_X \circ v_{\overline{X}}) \neq f(u)$ where X is the union of the influential sets that it has found thus far. If $f(u_X \circ v_{\overline{X}}) \neq f(u)$ then the binary search finds a new influential set.

In the binary search for a new influential set, the procedure defines a set X' that is equal to the union of half of the sets in $\{X_i\}_{i \notin I}$. Then either $f(u_{X \cup X'} \circ v_{\overline{X'}}) \neq f(u)$ or $f(u_{X \cup X'} \circ v_{\overline{X'}}) \neq f(u_X \circ v_{\overline{X}})$. Then it recursively does the above until it finds a new influential set X_{ℓ} .

It is easy to see that if f is s-term DNF then, whp, for all the terms T in f of size at least $c^2 \log(s/\epsilon)$, for all the random uniform strings u, v chosen in the algorithm and for all the strings z generated in the binary search, $T(u_X \circ v_{\overline{X}}) = T(u) = T(z) = 0$. Therefore, when f is s-term DNF, the procedure, whp, runs as if there are no terms of size greater than $c^2 \log(s/\epsilon)$ in f. This shows that, whp, each influential set that the procedure finds contains at least one variable that belongs to a term of size at most $c^2 \log(s/\epsilon)$ in f. Therefore, if f is s-term DNF, the procedure, whp, does not generate more than $c^2s \log(s/\epsilon)$ influential sets. If the procedure finds more than $c^2s \log(s/\epsilon)$ influential sets then, whp, f is not s-term DNF and therefore it rejects.

Let R be the set of all the variables that belong to the terms in f of size at most $c^2 \log(s/\epsilon)$. The procedure returns $h(x) = f(x_X \circ w_{\overline{X}})$ for random uniform w where X is the union of the influential sets $X = \bigcup_{i \in I} X_i$ that is found by the procedure. If f is s-term DNF then since $r = 4c^2(c+1)s\log(s/\epsilon)$ and the number of influential sets is at most $c^2s\log(s/\epsilon)$, whp, at least $(1/2)c\log(s/\epsilon)$ variables in each term of f that contains at least $c\log(s/\epsilon)$ variables not in R falls outside X in the partition of [n]. Therefore, for random uniform w, whp, terms T in f that contains at least $c\log(s/\epsilon)$ variables not in R satisfies $T(x_X \circ w_{\overline{X}}) = 0$ and therefore, whp, are vanished in $H = f(x_X \circ w_{\overline{X}})$. Thus, whp, H contains all the terms that contains variables in R and at most $cs\log(s/\epsilon)$ variables not in R. Therefore, whp, H contains at most $c(c+1)s\log(s/\epsilon)$ influential variables. From this, and using similar arguments as for the procedure **ApproxTarget** in the previous subsection, we prove that, **Approx**C makes at most $\tilde{O}(s/\epsilon)$ queries and

1. If f is s-term DNF then, whp, the procedure outputs X and w such that

 $= H = f(x_X \circ w_{\overline{X}})$ is s-term DNF⁸.

The number of influential variables in $H = f(x_X \circ w_{\overline{X}})$ is at most $O(s \log(s/\epsilon))$.

 $^{^{8}}$ So in this case we need the class to be closed under zero-one projection.

5:20 Optimal Testers

2. If f is ϵ -far from every s-term DNF then the procedure either rejects or outputs X and w such that, whp, $H = f(x_X \circ w_{\overline{X}})$ is $(3\epsilon/4)$ -far from every s-term DNF.

We can now run **Tester**C (with $3\epsilon/4$) on H from the previous subsection for testing C^* where C^* is the set of s-term DNF with $k = O(s \log(s/\epsilon))$ influential variables.

Algorithm 6 A procedure that removes variables from f that only appear in large size terms.

Algorithm $\operatorname{Approx} C(f, \epsilon)$

Input: Oracle that accesses a Boolean function f and ϵ Output: Either " $X \subseteq [n], w \in \{0, 1\}^n$ " or "reject"

Partition [n] into r sets

- 1. Set $r = 8sc \log(s/\epsilon)$.
- 2. Choose uniformly at random a partition X_1, X_2, \ldots, X_r of [n]

Find a close function and influential sets

- 3. Set $X = \emptyset$; $I = \emptyset$; t(X) = 0; k = 3ms
- 4. Repeat $M = 400k \ln(100k)/\epsilon$ times
- 5. Choose w and u uniformly at random from $\{0,1\}^n$;.
- $6. t(X) \leftarrow t(X) + 1$
- 7. If $f(u_X \circ v_{\overline{X}}) \neq f(u)$ then
- 8. Binary Search to find a new influential set X_{ℓ} ; $X \leftarrow X \cup X_{\ell}$; $I \leftarrow I \cup \{\ell\}$.
- 9. If |I| > k then output "reject" and halt.
- 10. t(X) = 0.
- 11. If $t(X) = 400 \ln(100k)/\epsilon$ then
- 12. Sample w uniformly at random from $\{0,1\}^n$;
- 13. $\operatorname{Output}(X, w, H = f(x_X \circ w_{\overline{X}})).$