



Peters, T., Holmes, M. V., Richards, B., Palmer, T. M., Forgetta, V., Lindgren, C. M., Asselbergs, F. W., Nelson, C. P., Samani, N. J., McCarthy, M. I., Mahajan, A., Davey Smith, G., Woodward, M., O'Keeffe, L. M., & Peters, S. A. (2021). Sex differences in the risk of coronary heart disease associated with type 2 diabetes: a Mendelian Randomization analysis. *Diabetes Care*, *44*(2), 556-562. [dc201137]. https://doi.org/10.2337/dc20-1137

Peer reviewed version

Link to published version (if available): 10.2337/dc20-1137

Link to publication record in Explore Bristol Research PDF-document

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# Sex differences in the risk of coronary heart disease associated with type 2 diabetes: a Mendelian Randomization analysis

Journal:	Diabetes Care
Manuscript ID	DC20-1137.R2
Manuscript Type:	Original Article: Cardiovascular and Metabolic Risk
Date Submitted by the Author:	n/a
Complete List of Authors:	Peters, Tricia; Lady Davis Institute for Medical Research; McGill University Faculty of Medicine, Division of Endocrinology, Department of Medicine, The Jewish General Hospital Holmes, Michael V; University of Oxford, National Institute for Health Research, Oxford Biomedical Research Centre; University of Oxford, Medical Research Council Population Health Research Unit; University of Oxford, Clinical Trial Service Unit & Epidemiological Studies Unit, Nuffield Department of Population Health; University of Bristol, Medical Research Council Integrative Epidemiology Unit Richards, Brent; McGill University, Medicine and Human Genetics; Lady Davis Institute for Medical Research, Centre for Clinical Epidemiology Palmer, Tom; University of Bristol, Bristol Medical School; University of Bristol, Medical Research Council Integrative Epidemiology Unit Forgetta, Vincenzo; Lady Davis Institute for Medical Research, Centre for Clinical Epidemiology Lindgren, Cecilia; University of Oxford, Wellcome trust centre for human genetics; Oxford University, Big Data Institute, Li Ka Shing Center for Health Information and Discovery; Broad Institute, Program in Medical and Population Genetics Asselbergs, Folkert; Utrecht University, Department of Cardiology, Division Heart & Lungs, University Medical Center Utrecht; Utrecht University, Institute of Cardiovascular Science, Faculty of Population Health Sciences; University of Leicester, Department of Cardiovascular Sciences; National Institute for Health Research Leicester Biomedical Research Centre, Glenfield Hospital Samani, Nilesh; University of Leicester, Department of Cardiovascular Science; National Institute for Health Research Leicester Biomedical Research Centre, Glenfield Hospital McCarthy, Mark; Oxford University, Wellcome Centre for Human Genetics; Oxford University Hospitals NHS Trust, Oxford National Institute for Health Research Biomedical Research Centre; Oxford University, Oxford Centre for Diabetes, Endocrinology and Metabolism Mahajan, Anubha; University of Oxford, O

Integrative Epidemiology Unit; University of Bristol, School of Social and Community Medicine
Woodward, Mark; George Institute, Professorial Unit; Oxford University, The George Institute for Global Health: Johns Honkins University
Department of Epidemiology
O'Keeffe, Linda; University College Cork, School of Public Health;
University of Bristol, Medical Research Council Integrative Epidemiology
Unit
Peters, Sanne; Utrecht University, Julius Center for Health Sciences and
Primary Care, University Medical Center Utrecht; University of Oxford,
George Institute for Global Health; George Institute for Global Health



# Sex differences in the risk of coronary heart disease associated with type 2

# diabetes: a Mendelian Randomization analysis

Sex differences in diabetes and heart disease

Tricia M. Peters, MD, PhD, Michael V. Holmes, MBBS, PhD, J. Brent Richards, MD, MSc, Tom Palmer, PhD, Vincenzo Forgetta, MSc, Cecilia M. Lindgren, PhD, Folkert W. Asselbergs, MD, PhD, Christopher P. Nelson, PhD, Nilesh J. Samani, MD, Mark I. McCarthy, MB, BChir, MD<sup>1</sup>, Anubha Mahajan, PhD<sup>1</sup>, George Davey Smith, MD, BChir, MSc, Mark Woodward, MSc, PhD, Linda M. O'Keeffe, PhD\*, Sanne A.E. Peters, PhD\* *\*Denotes equal contribution* 

# Affiliations

Centre for Clinical Epidemiology, Lady Davis Institute for Medical Research, Montreal, QC (T.M.P., J.B.R., V.F.)

Division of Endocrinology, Department of Medicine, The Jewish General Hospital, McGill University, Montreal, QC (T.M.P., J.B.R.)

Medical Research Council Population Health Research Unit, University of Oxford, Roosevelt Drive, Oxford, UK (M.V.H.)

Clinical Trial Service Unit & Epidemiological Studies Unit, Nuffield Department of Population Health, University of Oxford, Oxford, UK (M.V.H.)

National Institute for Health Research, Oxford Biomedical Research Centre, Oxford University Hospital, Oxford, UK (M.V.H.)

Medical Research Council Integrative Epidemiology Unit, University of Bristol, Oakfield House, Oakfield Grove, Bristol, UK (M.V.H., T.P., G.D.S., L.M.O.K.)

Population Health Sciences, Bristol Medical School, University of Bristol, Bristol, UK (T.P.) Big Data Institute, Li Ka Shing Center for Health Information and Discovery, Oxford University, Oxford, UK (C.M.L.)

Wellcome Centre for Human Genetics, Nuffield Department of Medicine, University of Oxford,

Oxford, UK (C.M.L, A.M., M.I.M.)

Program in Medical and Population Genetics, Broad Institute, Boston, MA, USA (C.M.L.)

Department of Cardiology, Division Heart & Lungs, University Medical Center Utrecht, Utrecht University, Utrecht, The Netherlands (F.W.A.)

Institute of Cardiovascular Science, Faculty of Population Health Sciences, University College London, London, United Kingdom (F.W.A.)

Health Data Research UK and Institute of Health Informatics, University College London, London, United Kingdom (F.W.A.)

Department of Cardiovascular Sciences, University of Leicester, Leicester, UK (C.P.N., N.J.S.) National Institute for Health Research Leicester Biomedical Research Centre, Glenfield Hospital, Leicester, UK (C.P.N., N.J.S.)

Oxford Centre for Diabetes, Endocrinology and Metabolism, University of Oxford, Churchill Hospital, University of Oxford, Oxford, UK (M.I.M., A.M.)

Oxford National Institute for Health Research Biomedical Research Centre, Oxford University Hospitals Trust, Oxford, UK (M.I.M.)

School of Social and Community Medicine, University of Bristol, Bristol, UK (G.D.S.) The George Institute for Global Health, University of Oxford, Oxford, UK (M.W., S.A.E.P.) The George Institute for Global Health, University of New South Wales, Sydney, Australia (M.W., S.A.E.P.)

Department of Epidemiology, Johns Hopkins University, Baltimore MD, USA (M.W.) School of Public Health, University College Cork, Ireland (L.M.O.K.) Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, Utrecht, The Netherlands (S.A.E.P.)

<sup>1</sup>Current address: Genentech, 1 DNA Way, South San Francisco, CA 94080

# **Contact for correspondence**:

Dr. Tricia M. Peters Centre for Clinical Epidemiology, Room H-450 Lady Davis Institute, Jewish General Hospital 3755 Cote Ste Catherine Montreal, Quebec H3T 1E2 CANADA Telephone number: 514-340-8222 x 28391 Fax number: 514-340-7529 Email address: tricia.peters@mcgill.ca

Total word count: 3689

# 3 Tables

1 Supplementary Figure, 3 Supplementary Tables

# Abstract

**Objective**: Observational studies have demonstrated that type 2 diabetes is a stronger risk factor for coronary heart disease (CHD) in women compared with men. However, it is not clear whether this reflects a sex differential in the causal effect of diabetes on CHD risk or results from sex-specific residual confounding.

**Methods**: Using 270 single nucleotide polymorphisms (SNPs) for type 2 diabetes identified in a type 2 diabetes genome-wide association study, we performed a sex-stratified Mendelian randomization (MR) study of type 2 diabetes and CHD using individual participant data in UK Biobank (N=251,420 women and 212,049 men). Weighted-median, MR Egger, MR-PRESSO and radial MR from summary-level analyses were used for pleiotropy assessment.

**Results**: MR analyses showed that genetic risk of type 2 diabetes increased the odds of CHD for women (odds ratio [OR] 1.13, 95% confidence interval [CI] 1.08-1.18 per 1-log unit increase in odds of type 2 diabetes) and men (OR 1.21, 95% CI 1.17-1.26 per 1-log unit increase in odds of type 2 diabetes). Sensitivity analyses showed some evidence of directional pleiotropy, however, results were similar after correction for outlier SNPs.

**Conclusions**: This MR analysis supports a causal effect of genetic liability to type 2 diabetes on risk of CHD that is not stronger for women than men. Assuming a lack of bias, these findings suggest that the prevention and management of type 2 diabetes for CHD risk reduction is of equal priority in both sexes.

# Introduction

Type 2 diabetes is a major risk factor for coronary heart disease (CHD)(1). Meta-analysis of observational studies demonstrates that type 2 diabetes is associated with a 44% greater relative risk of CHD in women compared with men(2). However, whether this reflects sex differences in the causal effect of type 2 diabetes on CHD or arises from confounding in observational studies is not well understood. Most observational studies adjust for traditional cardiovascular risk factors, yet novel biomarkers, social and behavioral factors, or women-specific risk factors, such as gestational diabetes, are not generally adjusted for and may explain some of the sex difference(3–5). Sex differences in screening for and treatment of type 2 diabetes might also contribute to the greater excess risk of CHD conferred by type 2 diabetes among women relative to men(6).

Mendelian randomization (MR) analysis exploits the natural random allocation of genetic variants at conception and is an increasingly utilized approach that can limit potential confounding in human research(7). Under the assumption that differences in the risk of disease arising from genotype mimic changes in the risk of disease acquired during life, MR can be used to detect causal effects. Recent MR studies support a causal relationship between genetic predisposition to type 2 diabetes and CHD(8,9). However, these studies did not evaluate sex differences in the causal role of type 2 diabetes in CHD risk. If type 2 diabetes has a stronger causal effect on CHD risk in women compared with men, randomly allocated genetic variants that are risk alleles for type 2 diabetes should also be more strongly associated with the risk of CHD in women than in men. Therefore, in this study we conducted a MR analysis to examine the sex-specific causal effect of the genetic risk of type 2 diabetes on CHD.

# Methods

# Data sources and study participants

Data from the UK Biobank and a consortium of genome-wide association studies (GWAS) for type 2 diabetes were used. The UK Biobank is a large prospective study of over 500,000 individuals(10). Baseline data collection in the UK Biobank was conducted between 2006 and 2010 across 22 assessment centers. Participants aged 37 to 73 completed touchscreen questionnaires, were interviewed by trained research nurses, had physical measurements taken and blood samples extracted and frozen. The presence of type 2 diabetes and CHD was self-reported at study baseline and confirmed by a trained nurse. Genotyping was performed using the Affymetrix UK BiLEVE Axiom array or the Affymetrix UK Biobank Axiom® array. A combined reference panel including UK10K samples was used for imputation(11). In accordance with the National Research Ethics Service and the governing Research Ethics Committee of UK Biobank, generic Research Tissue Bank approval was obtained, and study participants provided written informed consent(10).

For the present study, we included individual-participant data on 463,469 UK Biobank participants who had concordant genetic and self-reported sex, who clustered with the Great Britain population in 1000 Genomes(12), whose genetic data was of sufficient quality(13), and who provided data on type 2 diabetes and CHD at baseline. Individuals with self-reported type 1 diabetes, gestational diabetes only, or a diabetes diagnosis prior to the age of 18 were excluded. CHD was defined as self-reported history of angina or myocardial infarction, and linkage with hospital admissions data and the national death register was used to also identify incident diagnoses of CHD after the baseline visit using international classification of disease (ICD) 9 or 10 codes (ICD9 410-414, ICD10 I20-I25) using follow-up data from recruitment through the end of February 2016 (mean 5.3 [SD 2.4] years), with N=3453 incident cases of CHD for women and

N=7420 incident cases for men. Myocardial infarction was also defined using the UK Biobank algorithm (https://biobank.ctsu.ox.ac.uk/crystal/crystal/docs/alg\_outcome\_mi.pdf).

Sex-specific summary-level data (ß-coefficients and standard errors) for the genetic contribution of type 2 diabetes risk were obtained from the European DIAMANTE (DIAbetes Meta-Analysis of Trans-Ethnic association studies) GWAS of type 2 diabetes cases (N=30,053 women and 41,846 men) and controls (N=434,336 women and 383,767 men) of European descent(14). The UK Biobank was excluded from GWAS estimates used in our analyses to avoid sample overlap.

### Mendelian randomization and selection of SNPs for analyses

Mendelian randomization studies exploit the random assortment and independent inheritance of genetic variants in the population, which removes bias due to reverse causation and, if conducted appropriately, greatly reduces bias from residual or unmeasured confounding(15). However, three key assumptions must be met for genetic variants to serve as instrumental variables of an exposure in MR analyses (Supplemental Figure 1)(16). First, the variants must be associated with the exposure of interest; second, they must not be associated with confounders of the relationship between the exposure and the outcome; third, they must be independent of the outcome except for their association via the exposure. This third assumption relates to the issue of horizontal pleiotropy, in which one or more variants used in the instrumental variable influences the outcome via a pathway other than the exposure of interest. When horizontal pleiotropy has a net effect to bias the properties of the genetic instrument, the summary MR estimate can be biased either towards or away from the null. In this situation, horizontal pleiotropy leads to bias of the underlying 'true' causal effect and it is termed unbalanced horizontal, or directional, pleiotropy.

In this study, we used data from the UK Biobank for individual-participant MR analysis. SNPs with significant associations ( $p < 5x10^{-8}$ ) with type 2 diabetes from the sex-combined European DIAMANTE GWAS were selected (Supplemental Table 1). We assessed linkage disequilibrium (LD;  $r^{2}>0.2$ ) using PLINK(17) on a reference panel consisting of a random selection of 50,000 individuals from UK Biobank. Of 291 genome-wide significant SNPs from the European DIAMANTE GWAS, 270 were found in UK Biobank that were bi-allelic, were not in LD, and were not derived from GWAS that adjusted for body mass index. The SNPs were aligned to the same effect allele, and effect allele frequencies were checked for concordance. These 270 SNPs were used to generate sex-specific weighted genetic risk scores as the instrumental variable for analyses(18). Individual SNPs were coded as 0, 1, or 2 depending on the number of type 2 diabetes risk alleles. Each SNP was weighted by the corresponding sex-specific ß-coefficient obtained from the European DIAMANTE GWAS and then summed for all SNPs. This method reduces the risk of false positive results and bias toward the confounded observational association that may occur when all data (SNPs, exposure, outcome) are obtained from a single sample(19). Statistical analysis

The strength of the genetic risk score as an instrument for type 2 diabetes was assessed using the F-statistic, where an F-statistic greater than 10 provides evidence against the possibility of bias arising due to a weak instrument(20). The association of sex-specific genetic risk scores with potential confounders was evaluated to assess the validity of the second assumption of MR (i.e., the genetic instrument is not associated with potential confounders) and was also compared with the observational association of type 2 diabetes status with potential confounders.

Two-stage residual inclusion estimation using logistic regression at the second stage(21) and Terza standard errors(22) evaluated the association of the genetic risk scores for type 2

diabetes with CHD to estimate the odds of CHD per 1-log unit increase in the odds of type 2 diabetes. This method includes first-stage residuals to correct for endogeneity(21), since application of traditional instrumental variable estimation approaches can be problematic for models including a binary exposure and a binary outcome(23). Models were adjusted for age, genotype array, and the first four principal components of ancestry.

To assess and account for potential directional horizontal pleiotropy, we also performed summary-level MR analyses using SNP to type 2 diabetes estimates from DIAMANTE and SNP to CHD estimates in UK Biobank. For summary-level analyses, we obtained odds ratios (ORs) and 95% confidence intervals (CI) for the causal effect of a 1-log unit increase in the odds of genetic liability to type 2 diabetes on the odds of CHD using the weighted-median, MR Egger, Mendelian randomization pleiotropy residual sum and outlier (MR-PRESSO), and radial MR methods(24–27). The weighted-median method calculates a median of the SNP-specific causal estimates from the ratio method for each SNP(25). It has been shown to yield consistent estimates when the weights of up to half the instruments are not valid. The MR Egger method is equivalent to an inverse-variance weighted method but does not constrain the intercept to zero, and as such, the MR Egger estimate is the slope of the modified linear regression equation and the intercept represents the average pleiotropic effect across SNPs(24). A non-zero intercept provides evidence of unbalanced horizontal pleiotropy, and the slope of the regression coefficient should provide an estimate that is free from bias induced by unbalanced horizontal pleiotropy. Analyses were conducted using the 'MendelianRandomization' package in R Studio version 1.2.1206. The MR-PRESSO test detects and corrects for horizontal pleiotropy and was performed using the 'MRPRESSO' package in R(26). The first part of the test (MR-PRESSO global test) identifies the presence of horizontal pleiotropy, the second part corrects the causal estimate for identified

pleiotropy via outlier removal, and the third part (MR-PRESSO distortion test) tests whether the causal estimate significantly differs before and after correction. Additional analyses for pleiotropy assessment used radial MR Egger models to identify outliers in the UK Biobank analysis using the 'WSpiller/RadialMR' package in R with modified second order weights(27), and analyses were repeated after exclusion of sex-specific outliers. P-values for the test of interaction for estimates from separate analyses was used to assess interaction by sex for each analysis(28).

# Results

Characteristics of the UK Biobank participants are presented in Table 1 and Supplemental Table 3. The mean age was 57 (standard deviation [SD] = 8) years and 46% of participants were men. The prevalence of type 2 diabetes was 4% in women and 8% in men. CHD was documented among 5% of women (N=12 716) and 12% of men (N=26 344), with myocardial infarction diagnosed in 1.5% of women (N=3807) and 6% of men (N=12 871). Both women and men with CHD were more likely to have traditional CHD risk factors (older age, type 2 diabetes, history of smoking, dyslipidemia, and hypertension) (Supplemental Table 3).

The sex-specific 270-SNP genetic risk score showed a strong association with type 2 diabetes in both sexes (F-statistic 683 for women and 1005 for men, Supplemental Table 2), thus satisfying the first assumption of MR that the genetic instrument is associated with the exposure. We evaluated whether the apparent difference in instrument strength by sex was due to sex differences in the prevalence of type 2 diabetes. In a random subset of UK Biobank participants with 750 cases of type 2 diabetes for both women (N=18 493) and men (N=9100), the adjusted F statistic of 47 (R-squared 0.02) for women, and adjusted F statistic 45 (R-squared 0.03) for men were similar (data not shown). Thus, because the difference in instrument strength by sex is a

product of greater prevalence of type 2 diabetes in men, it is not likely to appreciably affect the comparative validity of estimates derived from MR analyses.

Potential confounders were similarly distributed across quartiles of the genetic risk score for both women and men (Table 2). Conversely, conventional observational analyses showed that type 2 diabetes status was strongly associated with all potential confounders assessed (Table 2), highlighting the need for instrumental variables in this setting.

Individual-participant results from TSRI analyses in UK Biobank showed similar effects of genetic risk of type 2 diabetes on CHD for each sex (OR 1.13, 95% CI 1.08-1.18 for women; OR 1.21, 95% CI 1.17-1.26 for men, Table 3). Sensitivity analyses using the weighted median method showed attenuated results (OR 1.04, 95% CI 1.00-1.08 for women; OR 1.06, 95% CI 1.03-1.09 for men, Table 3). Using MR Egger, evidence of directional pleiotropy was observed in women (OR 1.01, 95% CI 0.96-1.06 and intercept 0.004, 95% CI 0.000 to 0.008, Table 3) and men (OR 1.00, 95% CI 0.96-1.04 and intercept 0.008, 95% CI 0.004 to 0.011, Table 3). Results from MR-PRESSO after outlier correction were slightly attenuated compared with those from TSRI analyses for both women (three outliers removed, OR 1.08, 95% CI 1.05-1.13) and men (five outliers removed, OR 1.13, 95% CI 1.10-1.17, Table 3). Analyses excluding SNPs from the genetic instrument that were identified as outliers by radial MR showed similar effect estimates as the TSRI results: OR 1.09, 95% CI 1.05-1.14 for women; OR 1.24, 95% CI 1.20-1.29 for men (Table 3). We employed additional measures to assess for heterogeneity based on MR-Egger regression, including the Cochran Q-test and I-squared statistic. The Q-test showed evidence of heterogeneity in the effect of type 2 diabetes SNPs on CHD for both women (Q-statistic 395.8) and men (Qstatistic 666.0). The I-squared (I<sup>2</sup>) statistic measures heterogeneity in the genetic associations with the exposure, and results (I<sup>2</sup> 84.7% for women and 87.1% men) showed some evidence of

heterogeneity in the associations of SNPs with type 2 diabetes. Such heterogeneity could be reflective of multiple causal pathways between type 2 diabetes and risk of CHD.

# Discussion

In this MR study of the sex-specific effect of type 2 diabetes on CHD, we found that genetic predisposition to type 2 diabetes does not confer a greater excess risk of CHD for women than for men. While our results are consistent with previous sex-combined MR studies providing support for a causal role of type 2 diabetes in CHD risk(8,9), the finding that the causal effect of genetic liability to type 2 diabetes on CHD risk is not stronger for women than men is novel and differs from sex-specific estimates from the accumulated observational evidence(2). This includes a recent analysis in the UK Biobank, which showed a stronger association of type 2 diabetes with CHD for women than men(29).

There are several potential explanations for the differences between the findings of our MR study and the observational evidence. As with any observational study, studies of sex differences in the association of type 2 diabetes with CHD may not have controlled for all relevant confounders or may have controlled for confounders that were poorly measured, leading to residual confounding. If this residual confounding differs between the sexes, a sex difference in the observational association of type 2 diabetes with CHD could arise. For example, men are typically at higher absolute risk of CHD, and the prevalence of many cardiovascular risk factors is higher for men than for women(1). However, cardiovascular risk factors including type 2 diabetes appear to confer a greater relative CHD risk for women than for men in observational analyses(29). Furthermore, among individuals with type 2 diabetes compared to those without type 2 diabetes, several studies have shown that the differences in cardiovascular risk factors including blood pressure, dyslipidemia, and particularly anthropometric variables are greater among women than

men(3,6). Although women generally display a more favorable cardiometabolic risk profile than men, this favorable risk profile declines and ultimately reverses as glycemic control deteriorates(30).

Yet observational evidence of sex differences in the association of other major risk factors with CHD is not universally observed, suggesting mechanisms other than confounding alone may be involved. An alternative explanation is that sex differences in the effect of diabetes on CHD risk seen in observational studies reflect the more adverse deterioration in cardiovascular risk profile along the glucose intolerance spectrum in women than men. A recent MR study showed that the association of BMI with the risk of diabetes was stronger for women than men(31). Accordingly, a pathway of type 2 diabetes progression and glycemic dysregulation that leads to more adverse complications of diabetes for women than men may underpin the observational findings, rather than a direct sex difference in the effect of diabetes on CHD risk.

Furthermore, women may be perceived as having lower cardiovascular risk and consequently, type 2 diabetes and comorbid cardiovascular risk factors may be treated less aggressively(32,33). Guidelines for the diagnosis and treatment of type 2 diabetes and CHD are not sex-specific; our results of a similar causal association of type 2 diabetes with CHD by sex would support the notion that for a given state of glycemic dysregulation and burden of cardiovascular risk factors, prevention and management of type 2 diabetes for the reduction of CHD risk should be of equal priority for both women and men. In addition, sex-specific confounders, such as reproductive factors including gestational diabetes, are rarely adjusted for in observational studies that include both sexes; this could inflate the association of type 2 diabetes with CHD in women if the cumulative duration of the exposure to diabetes is greater, on average, among women than men. Sex-specific residual confounding may therefore explain some of the

discrepancy between the MR and observational evidence. Alternatively, the discrepancy might arise if the MR analysis does not account for genetic variation in the risk of type 2 diabetes that derives from sex chromosomes, as the GWAS data includes only autosomal SNPs. For example, a recent MR study observed a causal association of genetically determined testosterone (X chromosome) with increased type 2 diabetes risk for women but not for men(34). Multiple other mechanisms could also play a role in conferring higher CHD risks for women with type 2 diabetes compared with men independent of glucose dysregulation or diabetes, including sex differences in microvasculature such as vascular responsivity to aldosterone(35).

The diagnosis of type 2 diabetes is defined by a cut-point along a continuum of glycemia that is based on the risk of associated complications such as retinopathy(36). Accordingly, an individual with borderline glycemia who is not yet diagnosed with type 2 diabetes may display phenotypic and genetic similarity when compared to an individual with diagnosed diabetes. Exposure misclassification of this type would tend to bias individual-participant MR estimates toward the null, leading to underestimation of the MR results. In our individual-participant MR, this scenario would only affect our conclusion when pre-diabetes affected a differential proportion of women and men in the study population. Of note, this should not influence summary-level MR results as the exposure is fully defined by genotype.

There are several strengths of our study, including the use of MR, which under specific assumptions can be used to test the hypothesis that a particular risk factor is causal for an outcome(16). In accordance with the first assumption of MR, the sex-specific genetic risk scores were very strong instruments for type 2 diabetes for both women and men. Meeting the second and third assumptions of MR, the genetic risk scores were shown to be broadly independent of measured potential confounding factors. Furthermore, for both women and men, results of

sensitivity analyses after correction for outliers were similar to initial results. However, there are also limitations of our study. Although the genetic risk scores were strong instruments for type 2 diabetes, our instruments may have been underpowered to detect modest differences in sex-specific causal effects. Furthermore, our analysis used genetic risk scores derived from 270 genome-wide significant type 2 diabetes SNPs in the sex-combined European DIAMANTE GWAS(14). Genetic instruments obtained from the SNPs that are associated with type 2 diabetes in sex-specific GWAS could also have been constructed. However, the European DIAMANTE GWAS observed only one significant sex-differentiated SNP(14) and thus, the impact of the use of a sex-combined instrument is unlikely to have changed our results substantially. Moreover, such an instrument would not permit direct comparison of sex differences in the overall genetic predisposition to type 2 diabetes, but instead compares the causal effect of two distinct sex-specific instruments on CHD risk.

SNPs included in the genetic instruments for type 2 diabetes may affect CHD risk via pathways separate from their effect on type 2 diabetes risk, and these pathways could differ by sex. For example, there was some evidence of directional pleiotropy using MR Egger. However, the intercept for both men and women neared zero and MR-Egger generally lacks power. Moreover, results from outlier-robust sensitivity analyses were more similar to the overall results. This suggests that our primary results are in fact robust and that MR Egger results may have been influenced by sensitivity of this method to extreme outliers(37).

These results might reflect multiple different scenarios(38), some of which may have downstream effects on type 2 diabetes risk and may differentially affect CHD risk by sex. Taken together, we cannot exclude a sex-specific causal effect via other pathways not captured in our genetic instrument. Of note, our instrumental variables for type 2 diabetes were derived from the

DIAMANTE GWAS effect estimates without adjustment for BMI since the influence of BMI on type 2 diabetes risk may be sex-differential(31). Considering the important role of BMI in type 2 diabetes risk, adjusting for measures of adiposity in the type 2 diabetes GRS could bias a true differential effect of type 2 diabetes on CHD to the null. In addition, the UK Biobank and the European DIAMANTE GWAS used for our analyses include primarily European populations, and therefore, we cannot assess sex differences in the causal effect of type 2 diabetes with CHD across ethnicities. Furthermore, despite the large sample size of the UK Biobank, a low overall response rate of ~5.5% limits the generalizability of our results. Considering that the participating population is unlikely representative of the general UK population, as recently demonstrated(39), it is possible that our findings might be biased if there is a sex-specific selection bias that is associated with both the exposure and the outcome. Finally, a recent study demonstrated an association of autosomal loci with sex, which may introduce bias due to sex differences in study participation(40). If risk alleles for type 2 diabetes were associated with study participation in a sex-specific manner, this may have resulted in an inability to consistently detect a sex difference in the causal effect of type 2 diabetes with CHD in our MR analyses.

### Conclusion

The present MR analysis supports a causal effect of type 2 diabetes on the risk of CHD, with similar effects seen between women and men. In the absence of bias, these findings suggest that the prevention and management of type 2 diabetes for the reduction of CHD risk should be of equal priority for both women and men.

### Acknowledgments

We thank the European DIAMANTE investigators for making their data available. This research has been conducted using the UK Biobank Resource under Application Number '27449'.

Dr. Tricia Peters is the guarantor of this work and, as such, had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

# **Sources of Funding**

This research is supported by the Lady Davis Institute for Medical Research and the Department of Medicine, Jewish General Hospital (T.M.P.); the UK Medical Research Council, British Heart Foundation Intermediate Clinical Research Fellowship (FS/18/23/33512), and the National Institute for Health Research Oxford Biomedical Research Centre (M.V.H.); the Canadian Institutes of Health Research (CIHR), the Lady Davis Institute of the Jewish General Hospital, the Canadian Foundation for Innovation, the NIH Foundation, Cancer Research UK and the Fonds de Recherche Québec Santé (FRQS), and a FRQS Clinical Research Scholarship (J.B.R.); The Li Ka Shing Foundation, The National Institute for Health Research Biomedical Research Centre, Oxford, National Institutes of Health (CRR00070 CR00.01) and WT-SSI/John Fell funds, Widenlife and NIH (5P50HD028138-27) (C.M.L.); University College London Hospitals National Institute for Health Research Biomedical Research Centre (F.W.A.); the British Heart Foundation (C.P.N. and N.J.S.); Wellcome Trust 090532, 098381, 203141 and 212259, and NIDDK U01-DK105535, was a Wellcome Investigator and an NIHR Senior Investigator (M.I.M.); a UK Medical Research Council Population Health Scientist fellowship (MR/M014509/1) (L.M.O.K.); a UK Medical Research Council Skills Development Fellowship (MR/P014550/1) (S.A.E.P.).

# Disclosures

M.V.H. has collaborated with Boehringer Ingelheim in research, and in accordance with the policy of The Clinical Trial Service Unit and Epidemiological Studies Unit (University of Oxford), did not accept any personal payment. This study was supported by the NIHR Oxford Biomedical Research Centre. The views expressed are those of the author(s) and not necessarily those of the NHS, the NIHR or the Department of Health. J.B.R. has served as an advisor to GlaxoSmithKline. M.I.M has served on advisory panels for Pfizer, NovoNordisk and Zoe Global, has received honoraria from Merck, Pfizer, Novo Nordisk and Eli Lilly, and research funding from Abbvie, Astra Zeneca, Boehringer Ingelheim, Eli Lilly, Janssen, Merck, NovoNordisk, Pfizer, Roche, Sanofi Aventis, Servier, and Takeda. As of June 2019, M.I.M. is an employee of Genentech, and a holder of Roche stock. As of January 2020, A.M. is an employee of Genentech, and a holder of Roche stock.

## **Author Contributions**

T.M.P. analyzed the data, wrote the manuscript, and contributed to study design and conception. M.V.H. contributed to study design and reviewed/edited the manuscript. J.B.R. contributed to study design and reviewed/edited the manuscript. T.P. contributed to study design and data analysis and reviewed the manuscript. V.F. contributed to data analysis and reviewed the manuscript. C.M.L. reviewed/edited the manuscript. F.W.A. reviewed/edited the manuscript. C.P.N. reviewed/edited the manuscript. N.J.S. reviewed/edited the manuscript. M.I.M. contributed data and reviewed/edited the manuscript. A.M. contributed data and reviewed/edited the manuscript. G.D.S. reviewed/edited the manuscript. M.W. contributed to study design and reviewed/edited the manuscript. L.M.O.K. contributed to study design and wrote the manuscript. S.A.E.P. contributed to study design, data analysis, and wrote the manuscript.

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http://biorxiv.org/content/early/2020/03/23/2020.03.22.001453.abstract

	Women	(N=251 420)	Men (N=	212 049)
Age, mean $(SD^*)$ , years	56.6	(7.95)	57.0	(8.12)
Array type, No. (%)				
BiLEVE	24 920	(9.9)	24 897	(11.7)
Axiom	226 489	(90.1)	187 147	(88.3)
Type 2 diabetes, No. (%)	9964	(4.0)	16 917	(8.0)
Body mass index, mean (SD), kg/m <sup>2</sup>	27.0	(5.1)	27.9	(4.2)
Waist circumference, mean (SD), cm	84.6	(12.5)	97.1	(11.4)
Smoking history, No. (%)				
Never	146 521	(58.3)	102 139	(48.2)
Previous	81 252	(32.3)	82 970	(39.1)
Current	22 574	(9.0)	26 011	(12.3)
Dyslipidemia, No. (%)	25 549	(10.2)	33 843	(16.0)
Hypertension, No. (%)	57 721	(23.0)	64 668	(30.5)
Systolic BP <sup>†</sup> , mean (SD), mmHg	135.3	(19.1)	141.1	(17.4)
Diastolic BP, mean (SD), mmHg	80.5	(9.9)	84.0	(9.9)
Coronary heart disease, No. (%)	12 716	(5.1)	26 344	(12.4)
Myocardial infarction, No. (%)	3807	(1.5)	12 871	(6.0)
Angina, No. (%)	4864	(1.9)	10 219	(4.8)

 Table 1. Population Characteristics, UK Biobank (N=463 469)

\*SD: standard deviation; †BP: blood pressure

2203 (13.0)	23 931 (12.3)	Current smoking, N (%)	6555 (12.4)	6670 (12.6)	6492 (12.2)	6417 (12.1)	Current smoking, N (%)
106.3 (13.2)	96.2 (10.8)	Waist, mean (SD), cm	97.4 (11.2)	97.2 (11.3)	96.9 (11.4)	96.8 (11.5)	Waist, mean (SD), cm
31.2 (5.2)	27.6 (4.0)	Body mass index, mean (SD), kg/m <sup>2</sup>	28.0 (4.2)	27.9 (4.3)	27.8 (4.3)	27.7 (4.3)	Body mass index, mean (SD), kg/m <sup>2</sup>
95.4 (17.5)	85.4 (13.7)	Weight, mean (SD), kg	86.5 (14.2)	86.3 (14.3)	86.0 (14.3)	85.9 (14.4)	Weight, mean (SD), kg
174.7 (6.8)	175.9 (6.8)	Height, mean (SD), cm	175.7 (6.8)	175.8 (6.8)	175.8 (6.8)	176.0 (6.8)	Height, mean (SD), cm
16 917	195 132	No. participants	53 012	53 012	53 011	53 014	No. participants
			17.43-19.54	17.05-<17.43	16.67-<17.05	14.58-<16.67	Quartile range
Diabetes	No Diabetes		Q4	Q3	Q2	Q1	Men
1893 (19.0)	10 823 (4.5)	Coronary heart disease, N (%)	3348 (5.3)	3247 (5.2)	3077 (5.0)	3008 (4.8)	Coronary heart disease, N (%)
			4302 (6.8)	2560 (4.1)	1898 (3.0)	1204 (1.9)	Type 2 diabetes, N (%)
3525 (35.4)	22 024 (9.1)	Hypertension, N (%)	15 730 (25.0)	14 769 (23.5)	14 032 (22.3)	13 190 (21.0)	Hypertension, N (%)
5888 (59.1)	51 833 (21.5)	Dyslipidemia, N (%)	7148 (11.4)	6497 (10.3)	6108 (9.7)	5796 (9.2)	Dyslipidemia, N (%)
1058 (10.6)	21 585 (8.9)	Current smoking, N (%)	5781 (9.2)	5755 (9.2)	5564 (8.9)	5543 (8.8)	Current smoking, N (%)
99.1 (14.8)	84.0 (12.0)	Waist, mean (SD), cm	85.5 (12.7)	84.8 (12.5)	84.3 (12.4)	83.7 (12.3)	Waist, mean (SD), cm
32.5 (6.6)	26.8 (5.0)	Body mass index, mean (SD), kg/m <sup>2</sup>	27.3 (5.2)	27.1 (5.2)	27.0 (5.1)	26.7 (5.1)	Body mass index, mean (SD), kg/m <sup>2</sup>
84.6 (18.2)	70.9 (13.5)	Weight, mean (SD), kg	72.0 (14.1)	71.6 (14.0)	71.3 (13.9)	70.8 (13.8)	Weight, mean (SD), kg
161.4 (6.3)	162.7 (6.2)	Height, mean (SD), cm	162.5 (6.3)	162.5 (6.2)	162.6 (6.3)	162.8 (6.2)	Height, mean (SD <sup>†</sup> ), cm
9964	241 456	No. participants	62 855	62 855	62 854	62 856	No. participants
			16.68-<18.71	16.31-<16.68	15.94-<16.31	13.82-<15.94	Quartile range
Diabetes	No Diabetes		Q4	Q3	Q2	Q1	Women
'n	ational associatio	Observ		enetic risk score	Quartiles of go		
is	diabetes diagnosi	Type 2		2 diabetes risk	Genetic type		

Table 2. Association of sex-specific genetic risk scores (270 SNPs\*) for type 2 diabetes, by quartile, with potential confounders, and

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association of observational type 2 diabetes with potential confounders in the UK Biobank.

Coronary heart disease, 6136 (11.6) 6512 (12.3) 6663 (12.6) 7033 (13.3) Coronary heart disease, 21	Type 2 diabetes, N (%)         2157 (4.1)         3248 (6.1)         4495 (8.5)         7017 (13.2)         Type 2 diabetes, N (%)	Hypertension, N (%)         15 205 (28.7)         15 784 (29.8)         16 386 (30.9)         17 293 (32.6)         Hypertension, N (%)         53	Dyslipidemia, N (%)         7925 (14.9)         8349 (15.7)         8473 (16.0)         9096 (17.2)         Dyslipidemia, N (%)         27	
Coronary heart disease, 2 N (%)	Fype 2 diabetes, N (%)	Hypertension, N (%)	Dyslipidemia, N (%)	
21 132 (10.8)		53 919 (27.6)	27 627 (14.2)	
5212 (30.8)		6216 (36.7)	10 749 (63.5)	

\*SNP: single nucleotide polymorphism; †SD: standard deviation

Table 3. Mendelian randomization analysis of type 2 diabetes and risk of coronary heart disease, by sex, in UK Biobank<sup>\*</sup>. Results indicate the increased risk of coronary heart disease per 1-log unit increase in genetic risk of type 2 diabetes (odds ratio [OR] and 95% confidence interval [CI]).

		Women			Men	
	O	R (95% CI)	p-value	OR	(95% CI)	p-value
Two-stage residual inclusion estimation <sup>†</sup>	1.13	(1.08-1.18)	5.84 x 10 <sup>-08</sup>	1.21	(1.17-1.26)	2.31 x 10 <sup>-24</sup>
Weighted-median <sup>‡.</sup>	1.04	(1.00-1.08)	0.067	1.06	(1.03-1.09)	< 0.001
MR-Egger <sup>‡</sup>	1.01	(0.96-1.06)	0.81	1.00	(0.96-1.04)	0.99
MR PRESSO (outlier- corrected) <sup>‡</sup>	1.08	(1.05-1.13)	3.11 x 10 <sup>-05</sup>	1.13	(1.10-1.17)	1.57 x 10 <sup>-12</sup>
Sex-specific outliers removed <sup>†‡§</sup>	1.09	(1.05-1.14)	6.76 x 10 <sup>-05</sup>	1.24	(1.20-1.29)	2.78 x 10 <sup>-27</sup>
	Inter	cept (95% CI)	p-value	Intercept	(95% CI)	p-value
MR-Egger (intercept) <sup>‡</sup>	0.002	(0.000-0.008)	0.027	0.008	(0.004, 0.011)	<0.001
Q-test <sup>‡</sup>	395.8			666.0		
I-squared <sup>‡</sup>	84.7%			87.1%		

\*Genetic instrument comprised of N=270 SNPs for type 2 diabetes identified in European

# DIAMANTE GWAS.

<sup>†</sup>Results from two-stage residual inclusion estimation using individual participant data

and weighted genetic risk score in UK Biobank. Adjusted for age, genotype array,

principle components of ancestry. P-value for interaction=0.02.

<sup>‡</sup>Results from summary-level analyses using SNP-type 2 diabetes estimates from
DIAMANTE GWAS (excluding UK Biobank) and SNP-CHD estimates from UK
Biobank. P-values for interaction: weighted median 0.43; MR-Egger 0.76; MR-PRESSO
0.07.

<sup>§</sup>Analysis with type 2 diabetes genetic instrument comprised of N=258 SNPs for women and N=245 SNPs for men, after SNPs identified as sex-specific outliers using radial MR excluded from genetic instrument. P-value for interaction <0.001.



outcome except for their association via the exposure. SNP: single nucleotide polymorphism; C: Confounders be associated with confounders of the relationship between the exposure and the outcome; 3) The variants must be independent of the Assumptions of Mendelian randomization: 1) The variants must be associated with the exposure of interest; 2) The variants must not

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SNP*	Locus	Chr <sup>†</sup>	Effect	Non-	Beta	SE	Beta	SE	<b>OUTLIER</b> <sup>‡</sup>
			Allele	effect					
rs1005752	HMG20A	15	A	C	0.082	0.013	0.076	0.012	W,M
rs10096633	LPL	8	С	Т	980.0	0.019	0.056	0.017	W
rs10097617	TP53INP1	8	Т	С	0.05	0.012	0.043	0.011	
rs10193538	BNIPL	2	Т	G	0.025	0.012	0.033	0.011	
rs10195252	GRB14/COBLL1	2	T	C	0.066	0.012	0.036	0.011	
rs10228066	DGKB	7	T	C	0.058	0.012	0.061	0.011	
rs10406327	PEPD	19	C	G	0.01	0.012	0.043	0.011	
rs10406431	GIPR	19	A	G	0.041	0.012	0.06	0.011	W
rs1042725	HMGA2	12	Т	С	0.053	0.012	0.04	0.011	
rs1061810	HSD17B12	11	A	С	0.041	0.013	0.078	0.012	
rs10750397	ETSI	11	A	G	0.031	0.013	0.053	0.012	
rs10757283	CDKN2A/B	9	Т	C	0.017	0.012	0.029	0.011	
rs10830963	MTNRIB	11	G	C	0.1	0.013	0.1	0.013	
rs10842994	KLHDC5	12	C	T	0.079	0.015	0.089	0.014	
rs10882101	HHEX/IDE	10	T	C	0.1	0.012	0.11	0.011	
rs10893829	ETSI	11	Т	C	0.022	0.017	0.06	0.016	
rs10908278	HNFIB	17	Т	A	0.052	0.012	0.083	0.011	
rs10937721	WFSI	4	C	G	0.071	0.012	0.096	0.011	
rs10938398	GNPDA2	4	A	G	0.056	0.012	0.042	0.011	
rs10954772	PURG	8	T	C	0.037	0.013	0.026	0.012	
rs10962	HNFIB	17	C	G	0.021	0.015	0.053	0.014	
rs10974438	GLIS3	9	C	A	0.035	0.013	0.055	0.012	

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from the European DIAMANTE genome-wide association study. Outlier SNPs identified using radial MR. Mendelian randomization analyses. Beta coefficients and standard errors (SE) for the association of each SNP with type 2 diabetes Supplemental Table 1. Details of the 270 single nucleotide polymorphisms (SNP) used as a genetic instrument for type 2 diabetes in

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	0 016	0 042	0 017	0 051	Ţ	Þ	16	SPG7	rs12920022
	0.012	0.055	0.013	0.053	A	G	15	PRCI	rs12910825
	0.013	0.043	0.014	0.05	G	A	12	FBRSL1	rs12811407
	0.011	0.031	0.012	0.028	C	Ţ	8	BOPI	rs12719778
	0.011	0.023	0.012	0.042	G	C	8	TRHR	rs12680028
	0.013	0.035	0.013	0.054	A	C	4	LCORL	rs12640250
	0.012	0.062	0.012	0.047	Т	C	2	GCKR	rs1260326
	0.012	0.045	0.013	0.051	С	Ţ	18	BCL2A	rs12454712
	0.02	0.039	0.021	0.065	Ţ	G	-	PATJ	rs12140153
	0.011	0.039	0.012	0.043	C	G		DSTYK	rs12048743
	0.011	0.043	0.012	0.047	Ţ	C	9	UBAP2	rs12001437
W	0.011	0.039	0.013	0.035	C	G	6	VEGFA	rs11967262
	0.012	0.029	0.013	0.048	Т	C	ω	KIF9	rs11926707
	0.013	0.028	0.014	0.051	T	G	13	HMGB1	rs11842871
	0.038	0.15	0.04	0.1	C	T	11	CCND1	rs11820019
W	0.013	0.059	0.014	0.062	A	G	6	CENPW	rs11759026
	0.029	0.077	0.031	0.12	A	G	15	TCF12	rs117483894
	0.017	0.11	0.019	0.11	A	G	3	PPARG	rs11709077
	0.014	0.08	0.015	0.1	G	A	ω	ADCY5	rs11708067
	0.02	0.03	0.022	0.07	Ţ	C	22	YWHAH	rs117001013
	0.011	0.041	0.012	0.037	T	O	20	CEBPB	rs11699802
	0.013	0.036	0.014	0.072	C	G	2	GLI2	rs11688682
	0.018	0.056	0.019	0.049	G	A	2	FAM49A	rs11680058
	0.011	0.048	0.012	0.034	С	G	16	FAM57B	rs11642430
Μ	0.026	0.17	0.029	0.15	С	Т	5	PAM	rs115505614
	0.015	0.037	0.016	0.049	С	Т	7	FBXL13	rs11496066
	0.011	0.046	0.012	0.048	Т	С	1	PTGFRN	rs1127215
	0.013	0.09	0.014	0.08	С	Т	10	CDC123/CAMK1D	rs11257655
	0.011	0.059	0.012	0.026	G	C	6	MTND2P8	rs11137820
	0.012	0.044	0.013	0.045	G	A	15	LTK	rs11070332
Μ	0.015	0.057	0.016	0.046	Ţ	C	12	CCND2	rs11063028
	0.012	0.032	0.013	0.057	Т	G	11	INS/IGF2	rs11042596

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	0.012	0.054	0.013	0.064	A	G	10	NEUROG3	rs177045
	0.011	0.037	0.012	0.056	A	G	8	MSRA	rs17689007
Μ	0.013	0.046	0.014	0.033	С	G	18	WDR7	rs17684074
	0.011	0.045	0.012	0.031	G	Ţ	14	AKAP6	rs17522122
	0.011	0.041	0.012	0.038	C	Ţ	S	ITGAI	rs17261179
Μ	0.033	0.099	0.034	0.12	A	G	ν	ANKH	rs17250977
	0.014	0.07	0.015	0.052	C	Ţ	7	DGKB	rs17168486
	0.014	0.043	0.015	0.038	C	G	14	SLC7A7	rs17122772
	0.011	0.097	0.012	0.086	T	C	7	JAZFI	rs1708302
	0.031	0.12	0.033	0.13	A	G	ω	UBE2E2	rs17013314
	0.011	0.029	0.012	0.052	A	C	4	SLC9B1	rs1580278
	0.012	0.029	0.013	0.081	А	G	7	KLF14	rs1562396
	0.013	0.046	0.013	0.038	Т	С	8	PVTI	rs1561927
	0.026	0.095	0.028	0.11	G	Т	4	PCGF3	rs1531583
Μ	0.018	0.087	0.019	0.084	С	Т	1	NOTCH2	rs1493694
	0.058	0.18	0.061	0.27	G	A	8	CPQ	rs149364428
	0.056	0.3	0.061	0.11	С	Т	1	FAM63A	rs145904381
	0.029	0.1	0.032	0.15	Т	G	11	<i>QSER1</i>	rs145678014
	0.013	0.035	0.014	0.059	А	G	12	WSCD2	rs1426371
	0.011	0.11	0.012	0.13	Т	С	16	FTO	rs1421085
	0.039	0.085	0.04	0.11	С	A	11	PDE3B	rs141521721
	0.012	0.032	0.013	0.061	T	C	9	LINGO2	rs1412234
	0.11	0.22	0.11	0.25	G	A	10	TCF7L2	rs140242150
W	0.012	0.054	0.013	0.059	G	С	17	ZZEF1	rs1377807
	0.014	0.044	0.014	0.04	Т	G	15	PTPN9	rs13737
	0.013	0.085	0.013	0.057	А	G	13	SPRY2	rs1359790
	0.023	0.034	0.026	0.13	G	A	2	CYTIP	rs13426680
Μ	0.012	0.059	0.013	0.041	A	С	5	DMGDH	rs1316776
	0.023	0.058	0.025	0.05	Т	С	3	SHQ1	rs13085136
	0.018	0.063	0.019	0.088	Т	С	20	NKX2.2	rs13041756
	0.028	0.074	0.029	0.05	С	Т	2	GRB14/COBLL1	rs13024606
	0.011	0.034	0.012	0.02	С	A	4	PABPC4L	rs1296328

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	0.011	0.039	0.012	0.043	G	T	3	RFT1	rs2581787
	0.017	0.042	0.018	0.028	С	Т	15	ONECUTI	rs2456530
	0.011	0.026	0.012	0.024	G	A	S	ANKRD55	rs2431115
	0.011	0.068	0.012	0.044	G	A	2	BCL11A	rs243024
Μ	0.013	0.06	0.014	0.052	G	A	11	KCNQ1	rs231361
	0.019	0.053	0.02	0.078	C	H	11	KCNQ1	rs231349
	0.011	0.039	0.012	0.07	C	Ţ	S	POC5	rs2307111
	0.012	0.024	0.014	0.041	G	A	11	KCNQ1	rs2283220
	0.011	0.071	0.012	0.039	G	Ţ	10	PLEKHAI	rs2280141
	0.012	0.033	0.012	0.044	А	C	ω	ROBO2	rs2272163
Μ	0.012	0.051	0.013	0.029	G	A	20	RALY	rs2268078
	0.019	0.084	0.02	0.13	А	Ţ	12	HMGA2	rs2258238
W	0.012	0.057	0.013	0.046	G	A	2	CEP68	rs2249105
	0.012	0.037	0.013	0.04	Т	С	19	GIPR	rs2238689
	0.029	0.19	0.033	0.16	Т	С	11	KCNQ1	rs2237897
	0.012	0.098	0.013	0.068	А	C	11	KCNQ1	rs2237895
	0.011	0.02	0.012	0.031	С	Т	12	USP44	rs2197973
	0.012	0.011	0.013	0.038	А	G	4	USP46	rs2102278
	0.015	0.055	0.016	0.022	Т	G	12	CDKNIB	rs2066827
	0.011	0.059	0.012	0.045	G	O	2	CEP68	rs2028150
	0.011	0.05	0.012	0.032	C	G	4	FAM13A	rs1903002
	0.048	0.15	0.052	0.18	G	C	10	TCF7L2	rs184509201
	0.013	0.045	0.014	0.041	Т	O	22	PIM3	rs1801645
	0.029	0.11	0.032	0.17	С	Т	20	HNF4A	rs1800961
	0.033	0.18	0.036	0.16	С	Т	12	HNFIA	rs1800574
	0.011	0.062	0.012	0.049	С	G	12	TSPAN8/LGR5	rs1796330
	0.014	0.049	0.015	0.058	G	С	14	NRXN3	rs17836088
	0.014	0.036	0.015	0.076	С	Т	11	MAP3K11	rs1783541
	0.011	0.029	0.012	0.035	Т	G	3	PPARG	rs17819328
	0.013	0.043	0.014	0.037	Т	G	2	DTNB	rs17802463
	0.022	0.12	0.025	0.094	G	A	9	TLE4	rs17791513
	0.022	0.12	0.024	0.081	А	G	8	CASCII	rs17772814

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	0.014	0.084	0.015	0.086	G	T	1	MACF1	rs3768321
	0.013	0.029	0.015	0.058	С	Т	16	CLUAPI	rs3751837
	0.021	0.044	0.023	0.092	С	Ţ	4	HTT	rs362307
	0.015	0.019	0.016	0.047	С	Ţ	2	PABPC1P2	rs35999103
	0.015	0.042	0.016	0.06	Т	C	2	TMEM18	rs35913461
M	0.012	0.05	0.013	0.05	А	C	17	TTLL6	rs35895680
	0.014	0.052	0.015	0.056	C	Ţ	ω	UBE2E2	rs35352848
	0.016	0.069	0.017	0.039	G	A	12	KSR2	rs34965774
	0.013	0.044	0.014	0.017	G	A	10	TCF7L2	rs34855922
	0.013	0.035	0.013	0.047	G	C	17	MLX	rs34855406
	0.012	0.058	0.013	0.032	А	G	1	ABCB10	rs348330
M	0.018	0.066	0.02	0.043	Т	C	15	RASGRP1	rs34715063
M	0.013	0.068	0.014	0.025	G	A	13	RNF6	rs34584161
W	0.013	0.047	0.014	0.054	Т	A	20	TSHZ2	rs34454109
	0.012	0.024	0.013	0.035	C	Ţ	6	LRFN2	rs34298980
	0.011	0.07	0.012	0.049	Т	C	1	PROX1	rs340874
	0.011	0.041	0.012	0.024	G	A	S	PHF15	rs329122
	0.013	0.051	0.014	0.042	А	G	12	CCND2	rs3217860
	0.022	0.1	0.024	0.08	Т	C	12	CCND2	rs3217792
	0.011	0.025	0.012	0.053	G	A	19	FARSA	rs3111316
Μ	0.012	0.11	0.013	0.089	А	G	2	IRS1	rs2972144
	0.012	0.039	0.013	0.094	C	Ţ	16	CMIP	rs2925979
	0.013	0.045	0.014	0.058	А	G		GNG4	rs291367
	0.014	0.05	0.015	0.032	А	C	4	PDGFC	rs28819812
	0.011	0.029	0.012	0.03	С	A	3	ABCC5	rs2872246
	0.015	0.086	0.016	0.076	А	G	9	GPSMI	rs28505901
Μ	0.012	0.057	0.013	0.068	G	С	1	LYPLALI	rs2820446
	0.013	0.061	0.014	0.028	G	A	6	SOGA3	rs2800733
	0.012	0.03	0.013	0.049	A	C	5	ARL15	rs279744
	0.011	0.069	0.012	0.034	А	G	9	TLEI	rs2796441
	0.012	0.044	0.013	0.032	А	C	11	PDHX	rs2767036
	0.012	0.066	0.013	0.042	Т	G	10	NEUROG3	rs2642588

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	0.075	0.27	0.078	0.31	C	G	15	WDR72	rs528350911
	0.013	0.041	0.014	0.037	А	T	18	MC4R	rs523288
	0.011	0.058	0.012	0.083	Т	C	11	KCNJ11	rs5213
	0.012	0.044	0.013	0.066	Т	C	9	ABO	rs505922
	0.012	0.056	0.013	0.05	Т	С	8	BOPI	rs4977213
	0.012	0.036	0.013	0.038	A	G	6	BEND3	rs4946812
	0.013	0.078	0.013	0.052	C	Ţ	15	AP3S2	rs4932265
	0.012	0.056	0.013	0.066	G	A	11	INS/IGF2	rs4929965
	0.012	0.038	0.013	0.066	G	A	17	RAII	rs4925109
	0.018	0.072	0.019	0.092	C	Ţ	20	HNF4A	rs4810426
	0.012	0.04	0.013	0.045	G	A	19	MAP2K7	rs4804833
	0.012	0.058	0.013	0.05	Т	A	15	MAP2K5	rs4776970
W	0.011	0.032	0.012	0.043	G	A	6	SLC22A3	rs474513
	0.017	0.051	0.018	0.059	Т	С	9	QKI	rs4709746
	0.012	0.039	0.013	0.027	С	Т	3	RBM6	rs4688760
	0.011	0.047	0.012	0.058	Т	С	3	LPP	rs4686471
	0.013	0.069	0.014	0.068	С	Т	5	ANKRD55	rs465002
	0.012	0.067	0.013	0.065	A	G	δ	ZBED3	rs4457053
W, M	0.016	0.1	0.018	0.079	С	Т	19	TOMM40/APOE	rs429358
	0.011	0.029	0.012	0.028	А	G	16	FTO	rs4281707
	0.011	0.04	0.012	0.028	С	G	7	IGF2BP3	rs4279506
	0.014	0.059	0.015	0.062	Т	С	12	CCND2	rs4238013
	0.014	0.049	0.015	0.053	G	С	12	MPHOSPH9	rs4148856
Μ	0.011	0.024	0.012	0.033	С	Т	7	RELN	rs39328
	0.011	0.051	0.012	0.052	С	Т	3	ST6GAL1	rs3887925
	0.019	0.058	0.021	0.068	А	G	5	ANKH	rs3845281
Μ	0.015	0.054	0.016	0.028	А	G	5	ITGAI	rs3811978
Μ	0.012	0.05	0.013	0.042	G	A	19	ZC3H4	rs3810291
	0.012	0.11	0.013	0.1	A	G	8	SLC30A8	rs3802177
	0.014	0.036	0.015	0.057	А	С	9	TFAP2B	rs3798519
	0.015	0.052	0.017	0.046	A	G	3	PSMD6	rs3774723
	0.013	0.044	0.013	0.055	С	Т	2	RBMS1	rs3772071

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	0.02	0.055	0.021	0.051	G	С	5	ANKH	rs6885132
	0.011	0.045	0.012	0.05	A	G	s	MRPS30	rs6884702
Μ	0.011	0.044	0.012	0.048	G	A	4	SMARCADI	rs6821438
	0.014	0.049	0.015	0.059	C	T	11	ETSI	rs67232546
	0.011	0.037	0.012	0.032	G	A	2	THADA	rs6708643
	0.015	0.06	0.016	0.062	C	Ţ	16	ITFG3	rs6600191
	0.011	0.022	0.012	0.04	A	G	2	BNIPL	rs6545714
	0.021	0.09	0.023	0.096	A	G	22	MTMR3/ASCC2	rs6518681
	0.011	0.054	0.012	0.022	C	Ţ	ω	SLC12A8	rs649961
Μ	0.012	0.057	0.013	0.057	C	G	7	MNXI	rs6459733
	0.012	0.07	0.013	0.039	Т	C	6	VEGFA	rs6458354
	0.012	0.037	0.013	0.044	G	A	7	AOCI	rs62492368
	0.026	0.052	0.028	0.12	Т	A	3	TSC22D2	rs62271373
	0.029	0.071	0.032	0.13	С	Т	2	TMEM18	rs62107261
	0.018	0.053	0.018	0.06	Т	С	18	COMMD9	rs62080313
	0.012	0.031	0.013	0.035	Т	G	14	MARK3	rs62007683
	0.014	0.082	0.015	0.05	G	С	17	BPTF	rs61676547
	0.011	0.041	0.012	0.03	С	G	20	GNAS	rs6070625
	0.013	0.039	0.013	0.044	А	G	20	EYA2	rs6063048
	0.017	0.061	0.018	0.042	С	Т	17	ACE	rs60276348
	0.015	0.074	0.016	0.077	А	G	6	MHC	rs601945
	0.016	0.072	0.017	0.074	C	T	4	ACSLI	rs58730668
	0.018	0.061	0.019	0.041	Т	С	1	FAF1	rs58432198
	0.015	0.061	0.016	0.061	А	G	13	KL	rs576674
	0.012	0.028	0.013	0.04	G	A	22	EP300	rs5758223
	0.014	0.043	0.015	0.068	Т	A	8	XKR6	rs57327348
	0.012	0.046	0.013	0.038	Т	С	11	MTNRIB	rs57235767
Μ	0.012	0.085	0.013	0.069	С	G	12	HNFIA	rs56348580
	0.012	0.039	0.013	0.067	Т	C	4	MAEA	rs56337234
	0.013	0.048	0.013	0.042	C	Α	9	ZNF169	rs55653563
	0.099	0.32	0.1	0.32	G	C	11	INS/IGF2	rs555759341
	0.014	0.028	0.015	0.082	A	C	1	SEC16B	rs539515

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	0.011	0.021	0.012	0.0092	Т	C	10	TCF7L2	rs7918400
	0.011	0.041	0.012	0.029	Т	C	9	FOCAD	rs7867635
	0.025	0.038	0.028	0.041	G	A	16	FTO	rs78020297
	0.023	0.065	0.025	0.075	G	A	12	RMST	rs77864822
Μ	0.012	0.14	0.013	0.15	А	G	6	CDKAL1	rs7756992
	0.015	0.1	0.017	0.09	C	A	11	CENTD2/ARAP1	rs77464186
	0.013	0.046	0.014	0.034	А	G	S	RASAI	rs7719891
	0.052	0.51	0.056	0.42	G	Ţ	12	CCND2	rs76895963
	0.012	0.07	0.013	0.044	А	Ţ	4	TMEM154	rs7669833
	0.039	0.074	0.041	0.11	С	Ţ	S	ANKH	rs76549217
	0.025	0.066	0.026	0.011	G	A	3	ST6GAL1	rs7645517
	0.016	0.065	0.017	0.041	Т	A	ω	EGFEMIP	rs7629630
	0.028	0.11	0.029	0.047	G	Ţ	ω	CACNA2D3	rs76263492
	0.015	0.029	0.015	0.041	Т	O	19	INSR	rs75253922
	0.027	0.091	0.03	0.083	А	O	ω	MBNLI	rs74653713
	0.038	0.14	0.042	0.17	А	C	18	MC4R	rs74452128
	0.013	0.066	0.014	0.037	С	Т	22	PNPLA3	rs738408
	0.034	0.097	0.037	0.084	А	G	12	HNFIA	rs73226260
	0.02	0.087	0.021	0.085	А	C	18	TCF4	rs72926932
W	0.022	0.14	0.023	0.088	А	С	16	BCARI	rs72802342
	0.014	0.04	0.015	0.044	G	A	19	UHRF1	rs7249758
	0.012	0.025	0.012	0.045	Т	C	18	LAMAI	rs7240767
	0.012	0.03	0.013	0.044	G	С	17	GLP2R	rs7222481
	0.013	0.034	0.014	0.05	А	G	12	ITPR2	rs718314
	0.011	0.045	0.012	0.028	Т	С	15	USP3	rs7178762
	0.024	0.11	0.025	0.045	Т	С	17	NF1	rs71372253
	0.011	0.02	0.012	0.047	С	A	11	ELF1	rs7124681
	0.011	0.039	0.012	0.03	G	A	11	CRY2	rs7115753
	0.011	0.059	0.012	0.061	С	G	10	ZMIZI	rs703972
	0.012	0.054	0.013	0.056	G	A	5	ARL15	rs702634
	0.011	0.023	0.012	0.056	A	G	6	HAUS6	rs7022807
	0.014	0.03	0.015	0.033	С	A	7	CTTNBP2	rs6976111

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	0.015	0.037	0.016	0.036	A	G	18	GRP	rs9957145
M	0.013	0.095	0.013	0.056	A	G	ω	SLC2A2	rs9873618
Μ	0.012	0.066	0.013	0.067	G	A	ω	ADAMTS9	rs9860730
	0.019	0.071	0.021	0.057	G	C	3	TMCCI	rs9828772
	0.014	0.041	0.015	0.076	G	A	S	ANKRD55	rs9687832
	0.013	0.056	0.013	0.024	T	A	13	DLEUI	rs963740
	0.012	0.023	0.013	0.032	T	A	13	SRGAP2D	rs9563615
	0.013	0.03	0.013	0.03	Ţ	C	13	PCDH17	rs9537803
	0.014	0.054	0.015	0.032	Т	C	6	RREB1	rs9505097
	0.012	0.059	0.013	0.041	G	A	6	SLC35D3	rs9494624
	0.011	0.043	0.012	0.015	G	C	1	SRGAP2	rs9430095
	0.019	0.051	0.02	0.099	A	G	6	RREB1	rs9379084
	0.014	0.052	0.015	0.055	Т	С	7	CRHR2	rs917195
W	0.013	0.072	0.014	0.065	G	A	7	GCK	rs878521
	0.011	0.023	0.012	0.028	Т	С	16	NFAT5	rs862320
Μ	0.021	0.11	0.022	0.061	A	Т	19	TM6SF2	rs8107974
	0.012	0.028	0.013	0.033	A	G	16	ATP2A1	rs8046545
	0.011	0.051	0.012	0.046	С	G	15	C2CD4A/B	rs8037894
	0.013	0.026	0.014	0.058	Т	С	15	RASGRP1	rs8032939
	0.013	0.015	0.014	0.043	Т	G	14	CLEC14A	rs8017808
	0.02	0.16	0.021	0.08	Т	A	2	THADA	rs80147536
	0.012	0.035	0.013	0.042	A	G	14	SMEK1	rs8010382
	0.011	0.013	0.012	0.042	С	Т	13	IRS2	rs7987740
	0.033	0.17	0.035	0.15	G	O	1	PROX1	rs79687284

\*SNP: single nucleotide polymorphism <sup>†</sup>Chr: Chromosome <sup>‡</sup>Outliers identified using radial MR; W: women; M: men

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Supplemental Table 2. Association of sex-specific genetic risk scores for type 2 diabetes with

type 2 diabetes\*. Genetic risk score comprised of 270 SNPs from the European DIAMANTE

genome-wide association study.

	Women		Men	
F statistic	683		1077	
R-squared	0.02		0.03	
Odds ratio (95%	1.70	(1.66-1.73)	1.70	(1.67-1.73)
confidence interval) <sup>†</sup>				

\*Adjusted for age, genotype array, and four principal components of ancestry.

<sup>†</sup>Odds ratios for the risk of type 2 diabetes per standard deviation increase in type 2 diabetes genetic risk score.

Supplemental Table 3: Population Characteristics, UK Biobank (N=463 469), stratified by sex

and coronary heart disease status.

	Women (N=251 420)				Men (N=212 049)				
	Without CHD		With CHD		Without CHD		With CHD		
	(N=238 704)		(N=12 716)		(N=185 705)		(N=2	6 344)	
Age, mean (SD <sup>*</sup> ), years	56.3	(8.0)	61.6	(6.1)	56.3	(8.2)	61.5	(6.2)	
Array type, No. (%)									
BiLEVE	23 257	(9.7)	1663	(13.1)	21 256	(11.4)	3641	(13.8)	
Axiom	215 436	(90.3)	11 053	(86.9)	164 446	(88.9)	22 701	(86.2)	
Type 2 diabetes, No. (%)	8071	(3.3)	1893	(14.9)	11 705	(6.3)	5212	(19.8)	
Body mass index, mean	26.9	(5.1)	29.3	(5.8)	27.7	(4.2)	29.1	(4.6)	
$(SD), kg/m^2$									
Waist circumference,	84.2	(12.3)	90.8	(13.8)	96.5	(11.1)	100.9	(12.0)	
mean (SD), cm									
Smoking history, No. (%)									
Never	140 460	(58.8)	6061	(47.7)	93 156	(51.8)	8983	(34.1)	
Previous	76 325	(32.0)	4927	(38.7)	69 566	(37.5)	13 404	(50.9)	
Current	20 937	(8.8)	1637	(12.9)	22 237	(12.0)	3774	(14.3)	
Dyslipidemia, No. (%)	21 515	(9.0)	4034	(31.7)	24 773	(13.3)	9070	(34.4)	
Hypertension, No. (%)	51 335	(21.5)	6386	(50.2)	51 029	(27.5)	13 639	(51.8)	
Systolic BP <sup>†</sup> , mean (SD),	135.1	(19.1)	140.1	(19.7)	141.0	(17.2)	141.4	(18.7)	
mmHg									
Diastolic BP, mean (SD), mmHg	80.6	(9.9)	79.8	(10.5)	84.3	(9.8)	81.7	(10.6)	

\*SD: standard deviation; <sup>†</sup>BP: blood pressure