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3 **1 Developmental transitions in body color in chacma baboon infants:**
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6 **2 implications to estimate age and developmental pace**
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52 22 **ABSTRACT**

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54 23 **Objectives:** In many primates, one of the most noticeable morphological developmental traits
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56 24 is the transition from natal fur and skin color to adult coloration. Studying the chronology and
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58 25 average age at such color transitions can be an easy and non-invasive method to (i) estimate the
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26 age of infants whose dates of birth were not observed, and (ii) detect inter-individual differences
27 in the pace of development for infants with known birth dates.

28 **Materials and Methods:** Using a combination of photographs and field observations from 73
29 infant chacma baboons (*Papio ursinus*) of known ages, we (1) scored the skin color of six
30 different body parts from pink to grey, as well as the color of the fur from black to grey, (2)
31 validated our method of age estimation using photographic and field observations on an
32 independent subset of 22 infants with known date of birth, and (3) investigated ecological,
33 social and individual determinants of age-related variation in skin and fur color.

34 **Results:** Our results show that transitions in skin color can be used to age infant chacma
35 baboons less than seven months old with accuracy (median number of days between actual and
36 estimated age = 10, range=0-86). We also reveal that food availability during the mother's
37 pregnancy, but not during lactation, affects infant color-for-age and therefore acts as a predictor
38 of developmental pace.

39 **Discussion:** This study highlights the potential of monitoring within- and between-infant
40 variation in color to estimate age when age is unknown, and developmental pace when age is
41 known.

42
43 **Keywords:** *ageing method, primate, inter-individual variation, food availability*

45 INTRODUCTION

46 The relationship between age and physical maturation has been a major topic of interest in
47 behavioral and evolutionary ecology. Across species, different developmental paces shed light
48 on between-species life-history variations (Sibly, Grady, Venditti, & Brown, 2014). For
49 example, primates, including humans, have longer periods of lactation, associated with slower
50 postnatal growth and smaller litter size compared to other mammals of similar sizes (Dufour &

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3 51 Sauther, 2002). Across populations and individuals of a same species, variation in infant
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5 52 developmental pace can have important fitness consequences on later survival and future
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7 53 reproduction (Lindström, 1999) as well as on maternal reproductive success, via a trade-off
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9 54 balancing investment in current versus future progeny (Lee, Majluf, & Gordon, 1991;
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11 55 Maestripieri, 2002; Trivers, 1972).

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14 56 Developmental pace may vary in relation to various environmental, social, maternal and
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16 57 individual traits. Specifically, and as for all mammals, the quantity and quality of milk represent
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18 58 a primary influence on infant growth trajectory (Lee, 1996). For example, a positive impact of
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20 59 exclusive breastfeeding in the first 3-5 months of life on growth rates is observed among
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22 60 children in Bangladesh (Arifeen, Black, Caul, Antelman, & Baqui, 2001). In non-human
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24 61 primates, food availability between conception and weaning is an important source of variation
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26 62 in developmental pace, which is likely mediated by maternal condition for young infants who
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28 63 do not yet feed independently (Lee, 1996). As a result, the seasonal timing of birth, as well as
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30 64 maternal traits that may affect the nutritional status of mothers, may affect growth and
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32 65 development (Lee, 1996; Lummaa, 2003; Vandeleest & Capitanio, 2012). For example, in
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34 66 hierarchical societies, dominant females often have a better access to food, which may lead to
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36 67 higher reproductive rates and faster offspring development (Clutton-Brock & Huchard, 2013;
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38 68 Stockley & Bro-Jørgensen, 2011). Individual traits, like infant sex, may also impact the
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40 69 developmental pace of sexually dimorphic primates (Lonsdorf, 2017). Finally, selective
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42 70 pressures such as predation risk (Coslovsky & Richner, 2011; Fontaine & Martin, 2006),
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44 71 infanticide risk (Bădescu et al., 2016) or group size (Borries, Larney, Lu, Ossi, & Koenig, 2008)
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46 72 may further impact developmental trajectories, depending on the species considered.

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49 73 In more than 30% of primate species, one of the most noticeable and widespread
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51 74 morphological developmental traits is the transition from natal fur and skin color to adult
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53 75 coloration (Ross & Regan, 2000; Treves, 1997). Neonatal fur coloration may have evolved to
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3 76 promote allomaternal care, infant attraction and protection by other group members, or in
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5 77 response to infanticide risk (Ross & Regan, 2000; Treves, 1997). Even if the ultimate function
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7 78 of neonatal fur color remains unclear (Hrdy, 1976; Treves, 1997), careful monitoring of such
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10 79 developmental transitions at the individual level could have at least two important applications.
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12 80 First, if the age at which such transitions occur does not vary excessively across individuals, it
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14 81 can be an easy and non-invasive method to estimate individual ages in transitional infants for
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16 82 whom the date of birth is unknown (Altmann & Altmann, 1981). Such a method relies on the
17
18 83 detailed description of age-related changes in body coloration for a cohort of infants of known
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21 84 ages. This approach, which can be applied using observational data only, has significant
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24 85 advantages over many other common methods, such as body weight or dental eruption patterns,
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26 86 which rely on invasive measurements (Hohn, 2009; Morris, 1972).

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28 87 Second, for infants of known ages, inter-individual differences in the age of color
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30 88 transitions may reflect broader individual variation in developmental pace, if they correlate with
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32 89 physical growth or other indicators of behavioral development, such as social, feeding and
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34 90 locomotor development, or the temporal dynamics of the mother-offspring relationship
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37 91 (Barthold, Fichtel, & Kappeler, 2009; Treves, 1997). Few studies have investigated such
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40 92 correlations. Across species, the loss of natal pelage of 138 primate species is partially
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42 93 correlated with the temporal dynamics of mother-offspring spatial proximity, i.e. this color
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44 94 change begins around the period when the infant spends a considerable amount of time (25%)
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46 95 out of the mother's arm-reach (Treves, 1997). Within species, the transition from natal to adult
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48 96 fur coincides with advanced physical independence, such as independent locomotion and
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51 97 feeding on solid food as is the case in red-fronted lemurs (*Eulemur fulvus rufus*) (Barthold et
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53 98 al., 2009). In such cases, recording infant color changes at particular ages may represent one
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56 99 easy way to quantify individual differences in developmental pace, and an early-life indicator
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58 100 of weaning age, which is a crucial life history trait affecting maternal fitness.

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3 101 Baboons are among the most studied of the non-human primates, but our ability to
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5 102 estimate infant age and our understanding of the determinants of infant developmental pace in
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7 103 this species are still limited. Little work has been done to estimate infant age based on color
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9 104 transitions since two seminal papers on yellow baboons (*Papio cynocephalus*) forty years ago
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11 105 (Altmann & Altmann 1981; Rasmussen 1979) - although a short preliminary study has also
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13 106 been published on chacma baboons (*Papio ursinus*) (Whitehead, Henzi, & Piper, 1990). These
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15 107 studies report that the fur and skin of infants of both species show different coloration compared
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17 108 to adults. When they are born, the fur is black and the exposed skin on the face, ears, hands,
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19 109 and feet are pink, and they gradually acquire the physical traits of adults, grey skin and dark
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21 110 grey fur for chacma baboons (yellow fur for yellow baboons) in their first year of life. This
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23 111 transition may operate following a specific chronology, with some parts of the body turning
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25 112 grey before others (Altmann & Altmann, 1981). However, those studies have only described
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27 113 the chronology of the color transitions, while the accuracy of a method based on such transitions
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29 114 to estimate infant ages has never been assessed.

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35 115 More is known about developmental pace in baboons, but few studies have focused on
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37 116 the specific pace of developmental transitions in coloration. High levels of individual variation
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39 117 can be seen across various indicators of development including changes in fur and skin color,
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41 118 as well as in growth rate and behavioral development, with some infants maturing faster than
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43 119 others (Altmann 1980; Altmann and Alberts 2005; Altmann and Altmann 1981; Johnson 2003).
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45 120 Food availability in the environment is a major factor influencing individual variation in growth
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47 121 rate (Altmann & Alberts, 2005). The timing of the birth in the annual cycle could thus impact
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49 122 infant developmental pace, as seasonal variation in food availability largely occur across
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51 123 tropical Africa (Feng, Porporato, & Rodriguez-Iturbe, 2013). Moreover, maternal traits are also
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53 124 known to affect baboon infant development: yellow baboon infants of dominant females have
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55 125 faster growth rates and reach sexual maturity earlier than those of subordinates (Alberts &
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3 126 Altmann, 1995; Altmann & Alberts, 2005; Charpentier, Tung, Altmann, & Alberts, 2008). In
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5 127 this same population, offspring of multiparous females are larger for their age than offspring of
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7 128 primiparous females (Altmann & Alberts, 2005). Overall, while the determinants of individual
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9 129 variation in growth trajectories have already been explored, less is known about color
10
11 130 transitions in baboons. Examining such variation may contribute to a better understanding of
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13 131 processes affecting growth and physical maturation, and changes in the color of certain body
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15 132 parts could be used as indicators of developmental pace in baboon species.

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17 133 Using the skin and fur coloration of wild infant chacma baboons, we first characterize
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19 134 in detail the temporal sequence of physical maturation from natal to adult-like coloration to
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21 135 compare the chronology of these coloration transitions with other baboon populations and
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23 136 species. Second, we assess the accuracy of the developed method by using these dates of
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25 137 coloration transitions to estimate the ages of infants, and comparing age estimates with their
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27 138 known birth dates. We then examine the main determinants of individual variation in body color
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29 139 transitions for a cohort of infants of known ages, focusing on the influence of ecological factors
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31 140 (food availability during pregnancy and early lactation, timing of infant birth in the annual
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33 141 cycle), maternal traits (rank and parity) and individual traits (sex) on the age at transitions.
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143 **Material and Methods**

144 *1 Study site*

145 Data were taken by different field observers each year between 2005 and 2019 from three
146 habituated groups of chacma baboons (J and L since 2005, and M, a fission of group J, since
147 2016) living in Tsaobis Nature Park (22°23S, 15°44'50E). Chacma baboons at the Tsaobis
148 Baboon Project live in a semi-arid region of Namibia, on the edge of the Namib Desert. Annual
149 rainfall is low and variable which leads to relatively low, and highly unpredictable food
150 availability (Cowlshaw & Davies, 1997). Nonetheless, rainfall occurs most years, and peaks

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3 151 between January and April, resulting in an annual food peak of variable intensity and duration
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5 152 from February to May. The groups were followed daily on foot from dawn to dusk, allowing
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7 153 observers to collect a variety of demographic, life-history, behavioral and group spatial data.
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10 154 Each group member was known individually through ear notches and identifying marks, and
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12 155 each new birth in a group was recorded. Reproductive states of adult females were recorded
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14 156 every day in the field, and they were classified as: (i) pregnant, where pregnancy was
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16 157 determined *post hoc* following infant birth, and encompassed the six months between the
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18 158 conceptive cycle and the birth; (ii) lactating, for the period from infant birth until the resumption
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20 159 of cycling; and (iii) cycling, including both swollen females (i.e., sexually receptive with a
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22 160 perineal swelling) and non-swollen females between two swelling phases. Groups were not
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24 161 followed year-round, and each year a field season of variable length was conducted
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26 162 (median=126 days, range: 57-240 days).
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33 164 **2 Ecological and individual traits**

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36 165 To investigate the determinants of individual variation in physical development, we considered
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38 166 both ecological and individual traits. We estimated food availability with the normalized
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40 167 difference vegetation index (NDVI). This measure has previously been used as a proxy of food
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42 168 availability in our population (Baniel, Cowlshaw, & Huchard, 2018) and other baboon
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44 169 populations (Zinner, Peí Aez, & Torkler, 2001). We extracted the mean NDVI per 16-day
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46 170 period on a 500 m resolution from 2004 until 2019 on the NASA website (MOD13A1 product)
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48 171 (Didan, Barreto Munoz, Solano, & Huete, 2015) within the home ranges of the three habituated
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50 172 groups. Using the GPS locations recorded every 30 minutes by the observers during daily group
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52 173 follows, between 2005 and 2019, we computed 100% isopleth home ranges for our three groups
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54 174 separately using kernel density estimates ('kernelUD' function, h set to 'href') implemented in
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56 175 the adehabitatHR package (Calenge, 2006). We then computed a daily NDVI value for each
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3 176 group on its given home range using a linear interpolation between two known NDVI values of
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5 177 16 days-intervals.
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8 178 The social rank of adult females was established each year for each group separately using
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10 179 *ad libitum* and focal observations of dyadic agonistic interactions (supplants, displacements,
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12 180 threats, chases and physical attacks). We computed a linear hierarchy among adult females
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14 181 using Matman 1.1.4 (Noldus Information Technology, 2013), and then calculated a relative
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16 182 rank for each female controlling for group sizes. This relative rank was computed using the
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18 183 formula $(N-r)/(N-1)$, where N was the number of adult females in the group in a given year, and
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20 184 r was the absolute rank of the female (ranging from 1 to N, 1 being the most dominant).
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22 185 Although adult females' ranks are relatively stable through time, individuals' relative ranks
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24 186 change as females enter and leave the hierarchy through maturation and death, respectively.
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26 187 Each female was assigned one relative rank per year, ranging from 0 (for the lowest ranking
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28 188 female) to 1 (for the highest-ranking), as hierarchy is inherited and stable for female baboons.
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30 189 For analyses of infant development, we considered the mother's rank during the year her infant
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32 190 was born.
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37 191 The mother's parity was assigned using long-term life-history data. Females were
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39 192 considered primiparous between the birth of their first and second infant, and multiparous after
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41 193 the birth of their second infant.
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45 195 **3 Infant color scoring**

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49 196 Infant coloration was scored using photographs and direct field observations. An observation,
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51 197 which could be either from a photograph or from the field, was a set of scores of different body
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53 198 parts on a given day for a given infant baboon. We used photographs opportunistically taken
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55 199 by field observers between 2005 and 2018 of infant baboons aged between 0-19 months old
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57 200 (N=255 pictures of 109 infants in total, median=2 pictures per individual, range: 1-11 pictures
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3 201 per individual). We discarded photographs when luminosity or quality were too poor to give a
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5 202 color score. In addition, in 2018 and 2019, infant baboons were scored every two weeks using
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7 203 direct field observations, until they finally turned completely grey. We obtained 158 field
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9 204 observations on 28 infants in this way, i.e. with no photograph (median=5.5 observations per
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11 205 individual, range: 1-10 observations per individual). Infant baboons were scored on
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13 206 photographs by one observer (Author 2), and in the field by one observer (Author 1). We
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15 207 determined the inter-individual consistency of scores by scoring independently the same set of
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17 208 60 pictures, and compared score similarity with the intra-class correlation coefficient for each
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19 209 body part (Koo & Li, 2016). To do so, we computed intra-class coefficient looking at the
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21 210 absolute agreement of a single-fixed rater with a two-way mixed models, using the ICC function
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23 211 of the ‘psych’ package in R (Revelle, 2020). The scoring reliability was excellent for each body
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25 212 part (mean \pm SD, ICC = 0.92 ± 0.06 , see also Table S1 of the Supporting Information). Thus,
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27 213 in total, we scored coloration for 134 infants between 2005 and 2019.

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33 214 In addition to determining the color of the fur covering the whole body, we scored the skin
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35 215 color of the following body parts: ears, eye contours, hands and feet, muzzle, muzzle tip, and
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37 216 ischial callosities (see Figure 1 for locations of these body parts). For each body part, we
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39 217 attributed a score between 1 and 5 following previous work (Altmann & Altmann, 1981;
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41 218 Rasmussen, 1979), such that 1 is fully pink, 2 is distinctively pinker than grey, 3 is a mixture
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43 219 of pink and grey, 4 is distinctively greyer than pink, and 5 is fully grey (Figure 2). For the fur,
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45 220 1 is fully black, 3 is a mixture of black to grey, and 5 is fully grey; we did not use scores 2 and
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47 221 4 because the color transition for the fur is less obvious than for other body parts (Figure 2).
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49 222 We scored all the body parts that were visible in each photograph. The data that support the
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51 223 findings of this study are openly available in Mendeley data:
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53 224 <http://dx.doi.org/10.17632/65vk9tck25.1>.

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3 226 *4 Accuracy of age estimations based on fur and skin coloration*
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6 227 To describe the physical transition in fur and skin coloration and further test whether it is a
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8 228 reliable method to estimate the age of an infant, we focused on scores from a sample of 73
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10 229 infants of known age, that is, whose date of birth was known with a maximum uncertainty of
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12 230 10 days. These dates of birth were determined from having been present on either the day of
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14 231 birth (51 infants, hereafter called Sample 1) or the day of conception (22 infants, called Sample
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16 232 2). The conception day was identified as the exact day of deturgescence of the sexual swelling
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18 233 in the conceptive cycle, which was the cycle followed by a pregnancy and by no other cycle.
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20 234 We then estimated the dates of birth of these 22 infants by adding 190 days to the conception
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22 235 date, based on the fact that the mean gestation length is 190 days at Tsaobis with little variability
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24 236 (range: 181-200 days, SD = 5, N = 13 pregnancies where both conception and birth were
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26 237 observed). We considered only baboons aged less than 8 months, as older baboons, from 8 to
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28 238 19 months old, were all fully grey. We thus restricted our dataset to this Cohort 1, comprising
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30 239 242 observations from 73 infants: 148 photographs on 59 infants and 94 field observations on
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32 240 16 infants (including 2 infants that were both photographed and observed in the field in 2018).
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34 241 Specifically, we obtained 238 scores for ears, 201 for eye contours, 238 for fur, 203 for hands
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36 242 and feet, 239 for muzzle, 91 for muzzle tip and 143 for ischial callosities.
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43 243 To determine the body parts that showed the most reliable timing in color transition to
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45 244 predict ages, we computed a correlation matrix using the Spearman's rank correlation
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47 245 coefficient between infant age (in months) and color scores for each body part of individuals in
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49 246 Cohort 1. We also examined pairwise correlation coefficients between the different coloration
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51 247 scores to determine whether some body parts provided redundant information because they
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53 248 changed color at the same age. It is important to note that the dataset used to generate these
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55 249 correlations is pseudoreplicated to some extent given that some individuals contribute multiple
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57 250 observations. These correlation coefficients are nevertheless useful to compare values across
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3 251 body parts, but should be interpreted with caution to evaluate the statistical significance of any
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5 252 single correlation.

7 253 To determine the accuracy of coloration scoring as an estimate of infants' ages, we
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10 254 computed the difference between the actual known age and the estimated age of another,
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12 255 independent subset of 70 observations on 22 infants. To do so, we divided our Cohort 1 in two
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14 256 sub-samples, namely Sample 1, including 172 observations of 51 infants with precisely known
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17 257 dates of birth (less than 10 days of uncertainty), and Sample 2, including 70 observations on 22
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19 258 infants (with exactly 10 days of uncertainty on their date of birth, since it had been inferred
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21 259 from conception date). We use Sample 1 as the reference for age estimation at color transitions,
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24 260 and Sample 2 to validate our method on a fully independent sample.

26 261 With Sample 1, we first conducted a principal component analysis (PCA) taking into
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28 262 account the color scores of the seven body parts considered, using the function 'PCA' of the
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30 263 'FactoMineR' package (Husson, Josse, Le, & Maintainer, 2020). We had a considerable
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33 264 number of missing values in our dataset as it was generally impossible to score all body parts
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35 265 from a single picture. To conduct a PCA with missing values, we performed imputations to
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37 266 complete the dataset using the 'imputPCA' function of the 'missMDA' package (Josse &
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40 267 Husson, 2016), which uses an iterative algorithm taking into account similarities between
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42 268 observations, as well as relationships between the scores of different body parts. The first
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44 269 dimension of the PCA (PC1) explained 86.6% of the variance of the color scores (versus 8.8%
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47 270 for the second dimension), and was the only PC retained for downstream analyses. We then
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49 271 investigated the relationship between PC1 and infant age (in days). To do so, we compared
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51 272 several linear models, all with PC1 as response variable, and containing as fixed effects a
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53 273 polynomial function of age modelled with variable degrees (from 1 to 8). We selected the best
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56 274 model of this subset as the one minimizing the Akaike Information Criterion (AIC), and a
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58 275 polynomial function of age of degree 4 was retained in our model.
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3 276 Next, we followed four steps to quantify the accuracy of age estimation from Sample 2.
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5 277 First, we estimated the missing values of the Sample 2 dataset on color scores using the method
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7 278 described above ('missMDA' package) (Josse & Husson, 2016). Second, we predicted the
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9 values of PC1 using the color scores of the seven body parts, using the 'predict.PCA' function
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11 279 of the 'FactoMineR' package, that took into account the PC1 values computed from Sample 1
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13 280 (Husson et al., 2020). Third, we used the linear model (that was computed on Sample 1) to
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15 281 calculate one age estimate for each of these predicted values of PC1, i.e. for each observation
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17 282 of Sample 2. More precisely, we used the 'uniroot' function in R to solve the equation of 4
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19 283 degrees linking each PC1 value with infant age. Fourth, we computed the absolute difference,
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21 284 in days, between the actual known age (\pm five days of uncertainty) and this estimated age to
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23 285 quantify the accuracy of our method of age estimation.
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28 287 In order to investigate the effect of the number of body parts that were scored on the
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30 288 accuracy of our age estimates, as well as which body part was the most informative in this
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32 289 context, we repeated this process with a variable number of body parts. From the seven body
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34 290 parts initially considered, we first removed the fur scores from Samples 1 and 2, as the scores
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36 291 of this body part had the lowest correlation with PC1. We then repeated this process by
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38 292 progressively removing one body part at a time, based on the relative values of the correlation
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40 293 between PC1 and the raw scores of that body part, removing first those parts for which scores
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42 294 were least correlated to PC1. Chronologically, we first removed fur, followed by hands and
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44 295 feet, ischial callosities, muzzle and eye contours. Our last round of analyses contained only
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46 296 scores from ears and muzzle tip.
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51 297 We also tested whether the 5-level scale for color scores produced more accurate age
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53 298 estimations than a simpler 3-level scale with the following levels: (1) pink (score 1 in our 5-
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55 299 level scale), (2) transitional (pooling scores 2, 3, 4 in our 5-level scale), (3) grey (score 5 in our
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57 300 5-level scale). Using only one score instead of three for transitional colors can facilitate data
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3 301 collection in the field, and be less subjective across observers. We then repeated the same
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5 302 processes described for Samples 1 and 2, to quantify the accuracy of the age estimation under
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8 303 a 3-level scale of color scores for different number of body parts considered.
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10 304 We also tested the effect of the number of observations per individual on the accuracy of
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12 305 age estimates. In Sample 2, there were 70 observations from 22 infants (range = 1-8, median=2
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14 306 observations per infant). For all infants with more than one observation, we computed one
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17 307 estimated birth date by averaging the different birth date estimates from each observation. We
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19 308 then computed the difference, in days, between the actual and estimated birth dates as an
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21 309 indicator of estimation accuracy to compare the accuracy of age estimates obtained from a
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23 310 variable number observations of a given infant.
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26 311 Moreover, we investigated the potential effects of the actual age of an infant on the accuracy
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28 312 of age estimations per observation, to test whether the accuracy of our age estimates may
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30 313 decrease for older juveniles. To do so, we ran a linear mixed model, with the accuracy of age
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32 314 estimates per observation (i.e. the absolute number of days between actual age and estimated
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34 315 age) as a response variable, the actual age as the only fixed effect, and infant identity as the
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37 316 only random effect (to control for repeated observations). For this analysis, we used all
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39 317 observations available from Sample 1, for all 7 body parts scored on a 5-level scale.
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42 318 Finally, we were interested in quantifying the effects of observation types, i.e. field or
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44 319 photographic observations, on the accuracy of age estimation per observation. To do so, we ran
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46 320 a linear mixed model, with the accuracy of age estimates per observation as a response variable,
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48 321 the observation type as the only fixed effect, and infant identity as the only random effect (to
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50 322 control for repeated observations). For this analysis, we used all observations available from
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52 323 Sample 1, for all 7 body parts scored on a 5-level scale to estimate ages of Sample 2 (Sample 2
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54 324 contained 38 field observations from 5 infants and 32 photographic observations from 17
55
56 325 infants).
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5 327 *5 Determinants of individual variation in the pace of color changes*
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8 328 To identify the main determinants of developmental pace, we used 242 observations from the
9
10 329 full sample of 73 infants with known ages (Cohort 1), and ran general additive mixed models
11
12 330 (GAMMs). We first computed a weighted mean color score for each observation, where the
13
14 331 score of each body part was given a different weight depending on its correlation with infant
15
16 332 age (see above). More precisely, we attributed less weight to body parts that were less correlated
17
18 333 with age by dividing their color score by their correlation coefficient, on the basis that the
19
20 334 developmental pace of these body parts was less variable across individuals for a given age,
21
22 335 and so presumably less affected by environmental and individual factors. We thus obtained a
23
24 336 general score of color development across all body parts for a given observation. Our response
25
26 337 variable was the weighted mean color score per observation (continuous variable).
27
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30
31 338 GAMMs allow to fit non-linear relationships between the response variable and one or
32
33 339 more fixed effect(s), using thin plate splines (Wood, 2003). We first determined the best model
34
35 340 describing the age effect on mean color scores, with the identity of infant set as random effect
36
37 341 to take into account the lack of independence between multiple observations of a same
38
39 342 individual. To do so, we compared GAMMs containing either (1) a simple, continuous effect
40
41 343 of age; (2) age modelled as a polynomial of degree 2; (3) a polynomial of degree 3; or (4) age
42
43 344 modelled using a thin plate regression spline. We selected the best model as the one minimizing
44
45 345 the Akaike Information Criterion (AIC). We found that the age modelled using a thin plate
46
47 346 regression spline was retained in all our models (See Table S2 in the Supporting Information
48
49 347 for model selection on the weighted mean color score response variable for an example). We
50
51 348 then ran our global GAMMs, with age modelled with a spline, infant identity as random effect,
52
53 349 and including the following additional fixed effects:
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3 350 - *Food availability during infant growth.* We expected infants whose gestation and birth
4
5 351 occurred during periods of high food availability to develop faster for their age (Altmann &
6
7 352 Altmann, 2005; Johnson, 2003). To test which time period of infant development (prenatal
8
9 353 versus postnatal) was most sensitive to variations in food availability, we investigated the
10
11 354 effects of both food availability during pregnancy (approximated by the mean NDVI daily
12
13 355 values between conception and birth date for a given infant) and early lactation (approximated
14
15 356 by the mean NDVI daily values between birth and observation date for a given infant). As these
16
17 357 two indicators of food availability were highly correlated (Pearson correlation = 0.39,
18
19 358 confidence interval: [0.28; 0.49], $t=6.58$, $P\text{-value}<10^{-4}$), leading to collinearity, we included
20
21 359 them separately in two different models.

22
23
24
25
26 360 - *Date of birth.* We expected infants born at particular times of the year to develop more
27
28 361 quickly. The NDVI is only a partial proxy of food availability, and other components of
29
30 362 seasonal variation that are not necessarily captured by NDVI (such as temperature, photoperiod,
31
32 363 or particular plant phenologies) may further affect developmental pace. As a given date in the
33
34 364 annual cycle is a circular variable (at least when considering environmental seasonality), we
35
36 365 used a sine fixed effect expressed as follows to introduce the infant's date of birth, converted
37
38 366 in radians, into our multivariate linear model: $\sin(\text{Infant's date of birth} + \varphi)$
39
40
41
42 367 Where φ is the phase value. We changed the phase value φ (to 0, $\pi/6$, $\pi/3$, $\pi/2$, $2*\pi/3$, $5*\pi/6$) to
43
44 368 account for potential phase shifts across the year, i.e. to consider all the months of the year as
45
46 369 possible birth months maximizing faster development. For example, a positive effect of the sine
47
48 370 phase of $\pi/6$ would maximize February, and a negative effect will maximize August. We ran
49
50 371 sequentially six different multivariate models (GAMMs with the six different phase values)
51
52 372 containing all other fixed effects, and we selected the best phase as the one minimizing the AIC
53
54 373 (which is $\pi/3$ for the weighted mean color score model).
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3 374 - *Mother's dominance rank*. We expected infants from dominant females to develop more
4
5 375 quickly, following studies on other aspects of baboon developmental pace (Altmann & Alberts,
6
7 376 2005; Cheney et al., 2004; S. Johnson, 2003).

9
10 377 - *Mother's parity*. We expected infants from primiparous mothers to develop more
11
12 378 slowly, following studies on other aspects of baboon developmental pace (Altmann & Alberts,
13
14 379 2005).

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16
17 380 - *Infant sex*. We did not expect any sex difference in the pace of early morphological
18
19 381 development, following studies on other aspects of baboon developmental pace (Altmann &
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21 382 Alberts, 2005; Altmann & Altmann, 1981; Johnson & Kapsalis, 1995). Nevertheless, we
22
23 383 included this variable to control for potential unexpected sex differences.

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26 384 - *Group identity*, to control for possible differences between social groups.

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28
29 385 We further considered each body part separately (seven models) to test whether the same
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31 386 determinants affected all body parts or only some of them, and to evaluate which body parts
32
33 387 were most sensitive to ecological, individual and maternal traits. Here, our response variables
34
35 388 were the scores of one given body part, and were coded as ordinal, with five categories for each
36
37 389 body part (except for the fur, with only three categories). We ran all the GAMMs using the
38
39 390 'gam' function of the 'mgcv' package (Wood, 2003), in R version 3.5.0 (R Core Team, 2018).
40
41 391 Ordinal response variables for each body part were specified with the family 'ocat' of the
42
43 392 'mgcv' package (Wood, 2003). We computed parameter estimates for each fixed effect, with
44
45 393 Wald statistic tests (X^2) and P-values associated. For smooth effects, we computed the effective
46
47 394 degrees of freedom, with the Wald statistic test and the P-value. We considered an effect to be
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49 395 significant when its P-value was < 0.05 . We were also interested in quantifying the importance
50
51 396 of inter-individual variation in developmental pace, and so tested the significance of the random
52
53 397 intercept, looking similarly at its P-value. For each model, we also checked graphically the
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55 398 normality of the residuals' distribution and the accuracy of the number of knots used for the
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3 399 age thin plate spine (this parameter constrains the ‘wiggleness’ of the smooth, i.e. the number
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5 400 of bow points in the fitted curve) (Wood, 2003) using the ‘gam.check’ function of the ‘mgcv’
6
7 401 package (Wood, 2019). Graphical representations were made using the ‘mgcViz’ package
8
9 402 (Fasiolo, Nedellec, Goude, & Wood, 2018).

403

404 **RESULTS**

405 **1 Characterization of fur and skin color changes in chacma baboon infants**

406 All results associated with the characterization of fur and skin color changes in relation to age
407 are summarized in Figure 3. All baboons were fully grey (i.e., at all body parts) after 8 months
408 of age, with similar fur and skin coloration as adults. The fur, muzzle or hands and feet were
409 the first body parts to turn grey, on average between 1 and 2 months of age. Hands and feet, as
410 well as the muzzle, started to transition around 1 month old until becoming fully grey at approx.
411 4 months of age. Ears started to turn grey after 2 months and were totally grey by 5 months.
412 The last body parts to change color were the eye contours, the ischial callosities, and the tip of
413 the muzzle, which started to transition at approx. 3 months of age, and finished after 5 months.
414 The qualitative chronology of the transition was apparent in the correlation matrix (Table 1):
415 hands and feet scores were highly correlated with muzzle scores, and eye contours scores were
416 highly correlated with muzzle tip scores. Nonetheless, the order in which body parts change
417 color was not always consistent across individuals. For some infants, we observed ears turning
418 grey before muzzles and hands, or ischial callosities turning grey before the muzzle tip and ears.
419 Overall, a chacma baboon infant with entirely pink skin and black fur was almost always
420 younger than 2 months old, as at least the hands and feet started to turn grey after this age. An
421 infant with entirely grey skin and fur was at least 4 months old, as no infant was observed with
422 a grey tip of the muzzle before this age. Finally, a transitional infant was aged between 1 and 6
423 months old.

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5 425 **2 Accuracy of age estimates based on color scores**
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8 426 The color scores were strongly correlated with the known ages of infants for all body parts,
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10 427 though correlation coefficients were slightly lower for the fur and the ischial callosities (Table
11
12 428 1). This means that the scores of these two body parts produce less precise age estimates,
13
14 429 probably due to higher inter-individual variation in the age of transition of these parts compared
15
16 430 to others. For example, a 3-month-old baboon could exhibit the full range of colors for ischial
17
18 431 callosities, from fully pink to fully grey; and a 1-month-old baboon could exhibit a fully black
19
20 432 or a fully grey fur (Figure 3).

21
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24 433 In order to test the accuracy of infant body coloration as a method of age estimation, we
25
26 434 computed a composite score (using PCA) from 172 observations of all body parts from 51
27
28 435 infants of known ages (births observed, Sample 1), and then predicted this composite score to
29
30 436 estimate ages on an independent sample of 70 observations from 22 infants of known ages
31
32 437 (Sample 2). We obtained a median difference of 12 days between the actual age and the
33
34 438 estimated age using 70 observations, 7 body parts and a 5-level scale of color scores (Table 2).
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36 439 We also investigated the effects of the number of body parts considered, and the number of
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38 440 levels of the color scale. Decreasing the number of body parts considered only slightly
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40 441 decreased the accuracy of age estimates (Table 2), and the same conclusion applied for using a
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42 442 3-level color scale instead of a 5-level one – though using a 3-level scale in combination with
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44 443 few body parts decreased accuracy more substantially (Table 2). Our best age estimations per
45
46 444 observation were computed when considering the skin across all 6 body parts but excluding the
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48 445 fur, and using a 5-level color scale (absolute number of days of between actual and estimated
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50 446 ages: median = 10.7, range = 0.1–86.4). We also investigated the effect of the number of
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52 447 observations per infant on the accuracy of age estimates: the more observations, the more
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54 448 accurate were the estimates (Figure S1, Supporting Information). Finally, the accuracy of age
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3 449 estimates significantly decreased with age ($X^2 = 10.0$, P-value = 1.54×10^{-3}) (Figure S2,
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5 450 Supporting Information), and were lower for photographic compared to field observations
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7 451 (Figure S3, Supporting information), albeit the latter effect did not reach significance ($X^2 =$
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9 452 3.63, P-value = 0.06).

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12 453

14 454 **3 Determinants of individual variations in the pace of color changes**

16
17 455 Finally, we investigated the determinants of inter-individual variations in age-related color
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19 456 scores, looking at ecological (food availability during pregnancy or during early-lactation, birth
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21 457 timing), maternal (parity, rank) and individual (sex) predictors using GAMMs controlling for
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23 458 age with a spline regression (Figure 4). The model explained 93% of the variation in the
24
25 459 weighted mean color scores. We detected significant inter-individual differences in mean color
26
27 460 scores (Table 3), and in most body parts (except for the muzzle tips, for which we had fewer
28
29 461 infants scored compared to other parts, see Table S3 in Supporting Information). Our index of
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31 462 food availability during pregnancy influenced the weighted mean color score (controlled for
32
33 463 age): infants turned grey earlier considering all their body parts together when food during
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35 464 pregnancy was abundant (Table 3, Figure 5). In contrast food availability during lactation did
36
37 465 not (Table S4, Supporting Information). We did not detect any effect of maternal (rank and
38
39 466 parity) nor individual (birth date and sex) traits on variation in infant mean color scores
40
41 467 controlled for age (Table 3). Looking at individual variation in color transitions for each body
42
43 468 part separately, we found a positive effect of food availability during pregnancy on the score of
44
45 469 ischial callosities only (Table S3, Supporting Information). We also found a positive effect of
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47 470 the timing of birth on the scores of muzzle tips (maximizing scores for infants born in
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49 471 December-January) but not for other body parts (Table S3, Supporting Information). We did
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51 472 not detect any effect of food availability during lactation, maternal traits, i.e. maternal rank and
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53 473 parity, on variation in infant color for any body part (Table S3, Supporting Information).

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6 475 ***DISCUSSION***

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8 476 Our purpose in this study was twofold: (1) to assess how well infant chacma baboons with
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10 477 unknown ages can be aged on the basis of their skin and fur color, and (2) to explore potential
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12 478 environmental, maternal, and individual predictors of variation in the pace of infant color
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14 479 development for infants with known ages. The age estimates obtained using infant color scores
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16 480 were generally accurate (median number of days between actual age and estimate age per
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18 481 observations = 10, range=0-86 days). All the tools required for estimating ages in a chacma
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20 482 baboon population are provided in this paper. Figures 1 and 2 can be used to score baboon
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22 483 coloration directly in the field (or using photographs), and the methods section details how the
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24 484 first dimension of a composite score (using PCA), derived from these color scores, can be
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26 485 computed to estimate ages. Various recommendations can be made to improve the accuracy of
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28 486 age estimates:

- 29 487 1. Use the color scores from all body parts except fur, as age estimates were most accurate
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31 488 with this configuration. However, decreasing the number of body parts considered does
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33 489 not decrease the accuracy of age estimations dramatically. For example, considering
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35 490 only two body parts (ears and muzzle tip) instead of six reduces the accuracy of age
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37 491 estimations by only 4.5 days. Nonetheless, such a result may vary across primate
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39 492 species, and one should establish the relevant body parts to be considered in different
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41 493 species.
- 42 494 2. Score the same individual repeatedly, as an infant with 8 observations had a birth date
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44 495 estimation error of 5 days in average, compared to more than 30 days for infants with
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46 496 only a single observation. Ideally, infants should be scored as early as possible (as
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48 497 accuracy decreases as infants age) and regularly until it turns fully grey.

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3 498 3. Prioritize field observations. Photographic evidence can be less accurate because color
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5 499 scores may be influenced by fluctuating light conditions. If necessary, photographic
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7 approaches should exploit multiple high-resolution photographs of the same individual
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9
10 501 on a given day at different times and angles to control for light variation.
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12 502 4. Consider using a 3-level color scale, which is slightly less accurate than a 5-level scale
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14 503 but also more practical, and may maximize inter-observer reliability. However, such
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16 choice depends on the trade-off between the accuracy targeted and the conditions of
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18 data collection (e.g., single versus multiple observers).
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22 506 In addition, this study opens the possibility of a similar application for machine learning where
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24 507 an algorithm could predict the age of an infant based on pictures, as long as a large set of high-
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26 508 quality pictures (see recommendation #3) of known-age individuals is available to initialize the
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28 algorithm. This relatively recent method is based on mathematical and statistical approaches
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30 through which computers can ‘learn’ from data to then make predictions. Such methods have
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32 been used in a wide range of field applications, and are becoming increasingly common (Al-
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34 Jarrah, Yoo, Muhaidat, Karagiannidis, & Taha, 2015).
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38 513 Changes in the pigmentation of infants’ ears and muzzles at Tsaobis follow a similar pattern
39
40 514 to that observed in another population of wild chacma baboon living in Mkuzi Game Reserve
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42 515 (South Africa), with an onset of grey coloration appearing around 1-2 months, and the full
43
44 516 transition completed around 4-6 months of age (Whitehead et al., 1990). Our detailed
45
46 517 characterization of the age at which fur and skin turn grey indicates that infant chacma baboons’
47
48 coloration develops faster than yellow baboons (Altmann & Altmann, 1981; Rasmussen, 1979).
49 518
50 Whilst chacma baboon infants have all turned grey by 8 months of age, most yellow baboon
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52 infants are still in the transition phase at this age, and only exhibit adult color around 12 to 14
53 520
54 months of age, depending on the population (Altmann & Altmann, 1981; Rasmussen, 1979). In
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56 addition, in contrast to yellow baboons, the order in which body parts change color was not
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3 523 always consistent across individuals in this study. Discrepancies observed in the age and
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5 524 chronology of transition between baboon species highlight that such patterns are species-
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7 525 specific, and that the methodology presented here should be developed and validated separately
8
9 526 for each species.

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11
12 527 Inter-individual variation in the rate of maturation was high for each body part considered
13
14 528 in this study, as found in yellow baboons (Altmann & Altmann, 1981; Rasmussen, 1979).
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16 529 Specifically, the change in fur color was a less accurate indicator of age, as fur turned grey
17
18 530 relatively early and showed marked individual variation. Failure to derive accurate age
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20 531 estimates remains non-negligible (we had a maximum of 86 days difference between actual and
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22 532 estimated age when relying on 6 body parts), due to the existence of outlier individuals with
23
24 533 very fast or slow development. This suggests that, whilst this method of age estimation is not
25
26 534 perfect, infant coloration can provide a simple and non-invasive proxy to detect individual
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28 535 variation in developmental pace on a semi-quantitative scale. Nevertheless, despite their utility
29
30 536 for tracking other developmental factors (Barthold et al., 2009; Treves, 1997), studies linking
31
32 537 age-related changes in color to other indicators of developmental pace are very rare in baboons.
33
34 538 Only one study on Amboseli yellow baboons recorded that late changes in paracallosal skin
35
36 539 color can be seen as a first sign of developmental abnormality, being followed by delayed
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38 540 maturation, locomotor disability and finally death (Altmann & Altmann, 1981).

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42 541 Regarding the determinants of individual variations in color-for-age considering all body
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44 542 parts together, we found that higher food availability during pregnancy, but not during early-
45
46 543 lactation, accelerated the transition towards adult coloration. In line with this, infant yellow
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48 544 baboons from a food-enhanced group grew more rapidly than individuals from wild-foraging
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50 545 groups (Altmann & Alberts, 2005). Similarly, in Phayre's leaf monkeys (*Trachypithecus*
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52 546 *phayrea*) infants in larger groups transitioned from natal to adult fur color later, suggesting that
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54 547 food competition affects infant color maturation and development (Borries et al., 2008). Here,
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3 548 we distinguished pre-natal from post-natal food availability, and our results emphasize the
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5 549 importance of maternal condition during pregnancy for infant post-natal color development. In
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8 550 capital breeders like baboons or humans, females can store energy to use it later, and conception
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10 551 likelihood generally peaks during periods of high food availability to increase the success of
11
12 552 their reproductive attempt (Brockman & van Schaik, 2005). Numerous human studies similarly
13
14 553 show that maternal nutritional status during pregnancy has a significant impact on infant birth
15
16 554 weight, early-life development, health and survival (Emery Thompson, 2013; Martorell &
17
18 555 Gonzalez-Cossio, 1987). This study, by highlighting the effect of prenatal food availability on
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20
21 556 inter-individual variations in color-for-age, suggests that similar effects could occur in baboons,
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23
24 557 and that color-for-age may thus be a reliable indicator of other developmental dimensions.

25
26 558 Studies on various primates including baboons have also shown that the infants of dominant
27
28 559 females exhibit faster development in terms of growth rate and age at weaning (Altmann &
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30 560 Alberts, 2005; Fedigan, 1983; Pusey, 2012). Yet, surprisingly, maternal rank did not affect
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32
33 561 individual variation in infant color in this study. Given that color scores reflect inter-individual
34
35 562 variation in maternal nutritional condition (see above), it is possible that rank does not capture
36
37 563 such differences at Tsaobis, perhaps because most food resources are not monopolizable, or
38
39 564 because subordinate females adopt foraging and social strategies that help to mitigate the costs
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41
42 565 of their low rank (Marshall, Carter, Ashford, Rowcliffe, & Cowlishaw, 2015; Sick et al., 2014).

43
44 566 Interestingly, looking at the color development of each body part separately, we found that
45
46 567 the only body part affected by prenatal food availability was the ischial callosities. This suggests
47
48 568 that the color scoring of different body parts could be used for different purposes; for example,
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50 569 ischial callosities are a better indicator of developmental pace than of age. Further investigations
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52 570 of the determinants of color changes at specific body parts showed that infants born in
53
54 571 December-January have greyer muzzle tips for their age than others. However, the muzzle tip
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56 572 was scored on the lowest number of infants (N = 16), and such a small sample size questions
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2
3 573 the robustness of this result that was not replicated using other body parts. Using a larger dataset
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5 574 will be necessary to reach conclusive results.
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7
8 575 Our study of the determinants of inter-individual variation of color-for-age is preliminary
9
10 576 for several reasons. First, we were not able to explore the impact of other factors that are known
11
12 577 to influence infant color transitions, like variation in social dynamics and infanticide risk
13
14 578 (Bădescu et al., 2016) that are probably of considerable evolutionary significance for young
15
16 579 chacma baboons (Cheney et al., 2004; Palombit, 2003). Second, further studies are needed to
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18 580 explore the correlation between age-related changes in coloration and developmental
19
20 581 milestones like weaning age for chacma baboons. Testing whether color development co-varies
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22 582 with other dimensions of development would indicate if it could be used as a reliable indicator
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24 583 of developmental pace that may ultimately be connected to fitness. In line with this suggestion,
25
26 584 the transition from natal to adult fur coloration of wild ursine colobus (*Colobus vellerosus*)
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28 585 infants has been used as a proxy of developmental pace to show that high infanticide risk
29
30 586 accelerates infant development (Bădescu et al., 2016). Finally, examining the potential effects
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32 587 of infant coloration on protective behavior from other group members, and on the level of
33
34 588 alloparenting care received wherever it is relevant (Brent, Teichroeb, & Sicotte, 2008; Ross &
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36 589 Regan, 2000) could be an interesting perspective to test other hypotheses proposed to account
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38 590 for the evolution of natal coat in primates.
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44 591 Our study demonstrates that monitoring the skin color transition of infant baboons is a non-
45
46 592 invasive and accurate method to estimate age up to 8 months old, with a median accuracy of 10
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48 593 days, which can decrease further if the same individual is scored repeatedly. Our study
49
50 594 additionally highlights inter-individual variability in the ages of color transitions. This suggests
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52 595 that scoring the change in infants' body coloration can also be used to detect individual variation
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54 596 of developmental pace, and certain body parts are more variable than others for a given age,
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56 597 and thus represent promising candidates in this respect. In this context, food availability during
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3 598 prenatal life affected infant color-for-age in our population, suggesting that maternal nutritional
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5 599 condition during pregnancy plays a central role in infant color development. Further research
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7
8 600 is necessary to determine whether such scores correlate with broader aspects of development
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10 601 (morphological, behavioral, and physiological), and which body parts are most useful.
11

12 602

15 603 **DATA AVAILABILITY STATEMENT**

17 604 The data that support the findings of this study are openly available in Mendeley data:
18
19
20 605 <http://dx.doi.org/10.17632/65vk9tck25.1>
21

22 606

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48 617

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3 778 **FIGURE LEGENDS**
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6 779 **Figure 1:** Picture of an infant chacma baboon showing the different body parts scored in this
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8 780 study.
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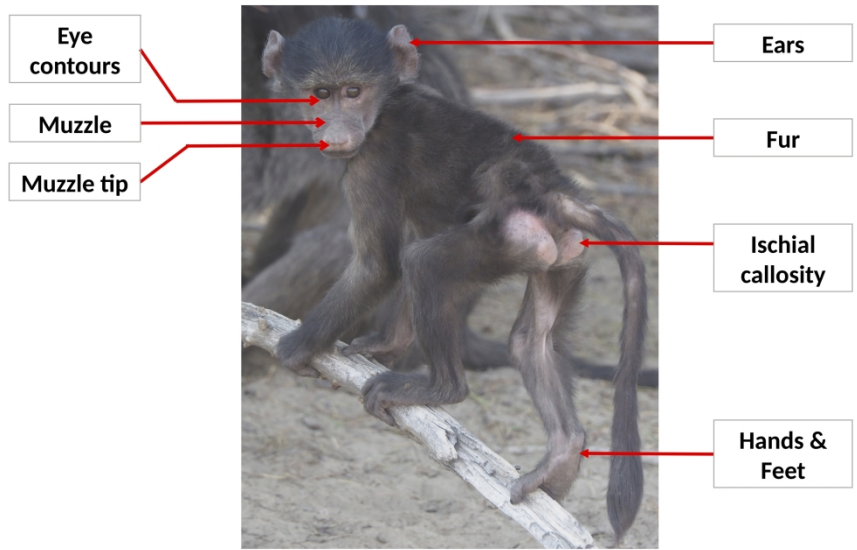
10 781 **Figure 2:** Pictures showing examples of color scores for different body parts. The first column
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12 782 refers to score 1, i.e. pink skin for each body part, and black fur. The second column refers to
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14 783 score 2, i.e., body parts that are pinker than grey (this score did not exist for fur). The third
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16 784 column refers to score 3, i.e., body parts that are just as pink as grey (and for fur, just as black
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18 785 as grey). The fourth column refers to score 4, i.e., body parts that are greyer than pink (this
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20 786 score did not exist for fur). The last column refers to the score 5, i.e. grey for all body parts as
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22 787 well as for fur.
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26 788 **Figure 3:** Color scores (from 1 to 5) of the different body parts of an infant according to its age
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28 789 (in months), using 242 observations on 73 infants with known birth date (Cohort 1). Month 0
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30 790 indicates an infant less than 1 month old, Month 1 indicates an infant aged between 1 and 2
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32 791 months old, etc. Each boxplot color represents a body part (see legend for details, ISC refers to
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34 792 ischial callosities). The median value for each boxplot is represented by a diamond, and outlier
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36 793 points are indicated by dark red dots.
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40 794 **Figure 4:** Weighted mean color score according to infant age (in days). Each dot represents the
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42 795 age component smooth function of the weighted mean color score according to individual age.
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44 796 The red curve shows the fitted smooth effect, and the 95% upper and lower confidence intervals
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46 797 are represented by the dashed blue curves.
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
































49 798 **Figure 5:** Effect of food availability during pregnancy on infant weighted mean color scores,
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51 799 controlled for infant age, identity and all other fixed effects included in our GAMM. Each dot
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53 800 represents an observation (raw values). The red curve shows the fitted linear effect of food
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55 801 availability, and the dashed blue curves represent the 95% upper and lower confidence intervals.
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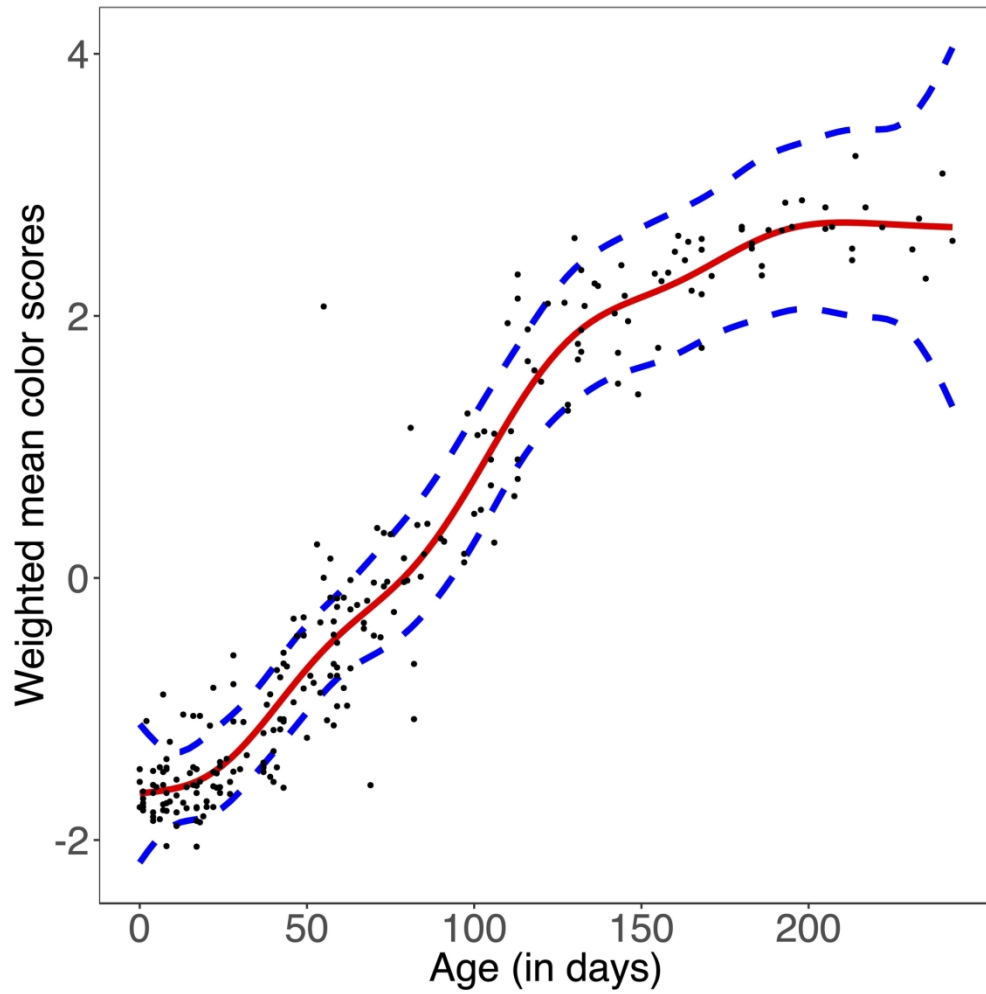


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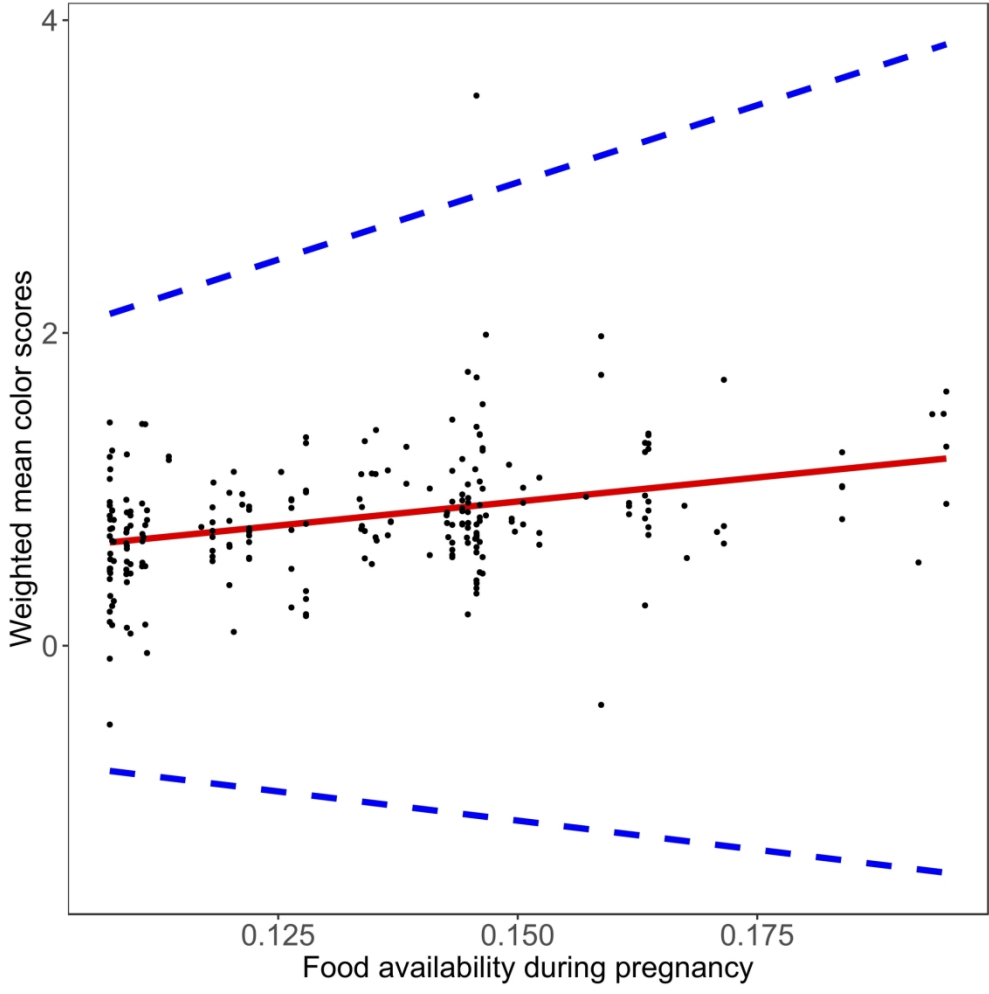
Score	1	2	3	4	5
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Ears					
Eye contours					
Hands & feet					
Muzzle					
Muzzle tip					
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