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1 The Single Component Geochemical Map: Fact or Fiction?

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241. Single component geochemical maps are the most basic representation of spatial elemental
25 distributions and commonly used in environmental and exploration geochemistry. However,
26 the compositional nature of geochemical data imposes several limitations on how the data
27 should be presented. The problems relate to the constant sum problem (closure), and the
28 inherently multivariate relative information conveyed by compositional data. Well known is,
29 for instance, the tendency of all heavy metals to show lower values in soils with significant
30 contributions of diluting elements (e.g., the quartz dilution effect); or the contrary effect,
31 apparent enrichment in many elements due to removal of potassium during weathering.
32 The validity of classical single component maps is thus investigated, and reasonable
33 alternatives that honour the compositional character of geochemical concentrations are
34 presented. The first recommended such method relies on knowledge-driven log-ratios,
35 chosen to highlight certain geochemical relations or to filter known artefacts (e.g. dilution
36 with SiO₂ or volatiles). This is similar to the classical normalisation approach to a single
37 element. The second approach uses the (so called) log-contrasts, that employ suitable
38 statistical methods (such as classification techniques, regression analysis, principal
39 component analysis, clustering of variables, etc.) to extract potentially interesting
40 geochemical summaries. The caution from this work is that if a compositional approach is
41 not used, it becomes difficult to guarantee that any identified pattern, trend or anomaly is
42 not an artefact of the constant sum constraint. In summary the authors recommend a chain
43 of enquiry that involves searching for the appropriate statistical method that can answer the
44 required geological or geochemical question whilst maintaining the integrity of the
45 compositional nature of the data. The required log-ratio transformations should be applied

- 46 followed by the chosen statistical method. Interpreting the results may require a closer
- 47 working relationship between statisticians, data analysts and geochemists.
- 48 Key words: soil geochemistry, compositional data analysis, log-ratios, mapping.

49 INTRODUCTION

50 Over the last decade the focus of digital soil and sediment databases at a regional, national,
51 transnational and continent scale has increasingly become to provide information for a
52 range of purposes including geological and soil mapping, baseline quality documentation,
53 mineral prospecting, land and soil resource assessment, risk evaluation, environmental and
54 educational purposes, and prediction of soil provenance for forensic purposes (Smith et al.,
55 1997; Reimann et al., 1998, 2003; Morris et al., 2003; Lagacherie et al., 2007; McKinley,
56 2013). These digital soil databases are generally accompanied by geochemical atlases (e.g.
57 Caritat and Cooper, 2011a; Young and Donald, 2013; Reimann et al., 2014a,b) showing a
58 collection of distribution maps for individual geochemical elements. This has been matched
59 by a corresponding increase in published studies utilising these soil geochemical surveys
60 (e.g. Chiprés et al., 2008; Grunsky et al., 2009; Carranza, 2010; Ohta et al., 2011; Caritat and
61 Grunsky, 2013; Cheng et al., 2014; Lancianese and Dinelli, 2014; Birke et al., 2015). The
62 resolution of the ground-based sampling scheme used for the generation of these databases
63 is the best compromise between the extent of the region covered, and time and resources
64 available. Over a local to regional scale, soil sampling can be managed on a 2 km² grid as
65 applied by the Geological Survey of Northern Ireland (GSNI)'s Tellus project. A regional scale
66 dataset such as this is close to exhaustive sampling in terms of geological survey mapping,
67 reducing the need for interpolation. It is worth noting however that this is not exhaustive
68 sampling for mineral exploration. Generally surveys are carried out at local (1:10,000),
69 regional (1:250,000) or continental (1:1,000,000) scales. A number of authors discuss the
70 scale of geochemical mapping (e.g. Bølviken et al., 1992; Darnley et al., 1995; Reimann et al.,
71 2010).

72 The sampling strategy generally follows standard protocols (e.g. UNESCO's IGCP Global
73 Geochemical Database – Darnley et al., 1995; G-BASE standard established by BGS - Johnson
74 2005; FOREGS field handbook – Salminen et al., 1998; GEMAS field handbook – EGS, 2008;
75 North American Soil Geochemical Landscapes Project – Smith et al. 2011; National
76 Geochemical Survey of Australia field manual – Lech et al., 2007; Canadian component of
77 the North American Soil Geochemical Landscapes Project – Friske et al., 2013; China
78 Geochemical Baseline Project – Wang et al., 2015).

79 Geochemical survey data are typically represented and interpreted using single element
80 geochemical maps. The interpretability and validity of these single components have
81 repeatedly been challenged because they are prone to several artefacts: spurious negative
82 bias on correlations (Chayes, 1960), dependence of interpretation on other (potentially non-
83 reported) components (Aitchison, 1986), dependence on units (e.g. mass, molar), and
84 dependence on processes acting on some components (e.g. weathering, dilution) but
85 influencing all of them (van den Boogaart and Tolosana-Delgado, 2013). All the issues
86 mentioned are due to the fact that geochemical data constitute amounts of components
87 with relative portions of a total even if this total is unknown. The components may be
88 reported in different physical units (ppm, mg/kg or as percentages) and all the components
89 may not be reported or measured. However, each component has an amount which
90 represents its importance as part of the whole composition. The constraints of constant sum
91 or the closed nature of the relative amounts of components have implications for the
92 analysis of geochemical data. In statistics and mathematical geosciences, powerful solutions
93 to deal with these issues have been developed in a field known as Compositional Data
94 Analysis (CoDA) (Aitchison, 1986; Grunsky, 2010; Pawlowsky-Glahn and Buccianti, 2011; van
95 den Boogaart and Tolosana-Delgado, 2013; Templ et al., 2011). This paper investigates the

96 question of what the compositional nature of geochemical data means in regional
97 geochemical mapping, specifically with regard to single element (univariate) distribution
98 maps.

99 The collaborative research presented stems from the first GeoMap Workshop (held in
100 Olomouc, Czech Republic, 17-20 June 2014) that discussed the consequences of these
101 challenges and the usefulness of CoDA for regional geochemistry. Present were
102 representatives from regional geochemical surveys (Tellus Survey covering Northern Ireland
103 and Tellus Border Survey covering the border counties of the Republic of Ireland (Young and
104 Donald, 2013)), the GEMAS project covering a large part of Europe (Reimann et al. 2014a,b),
105 the soil geochemical survey of the conterminous USA (Smith et al., 2011; Drew et al., 2010)
106 and the continental scale National Geochemical Survey of Australia (Caritat and Cooper,
107 2011a,b). This paper, while acknowledging the historically important role of single
108 component maps, aims to: (1) discuss their appropriateness, (2) provide some examples to
109 highlight the problems raised above, and (3) offer some solutions to present interpretable
110 maps free of the issues arising from the compositional nature of geochemical data.

111

112 PROBLEMS RELATED TO SINGLE COMPONENT MAPS

113 Geochemical surveys generate datasets with several tens of components (between 50 and
114 up to 70 elements are commonly reported), obtained from different sample materials (soil
115 horizons, size fractions, vegetal tissues, sediments, water, etc.) and with different analytical
116 techniques (total analyses, partial or selective digests). For regional geochemistry, the key
117 applications of the data are generally either to produce and use elemental concentration
118 maps (i.e. one-component regional distribution maps) or to explore associations between

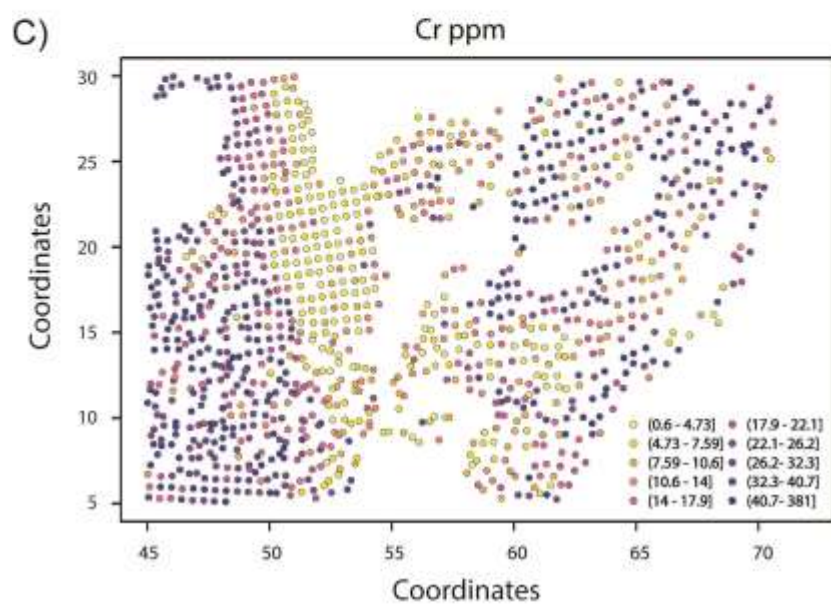
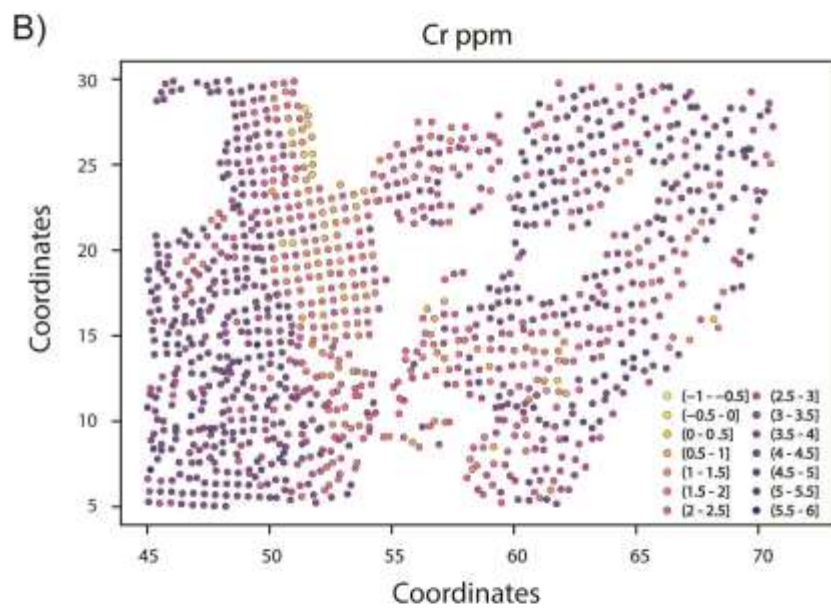
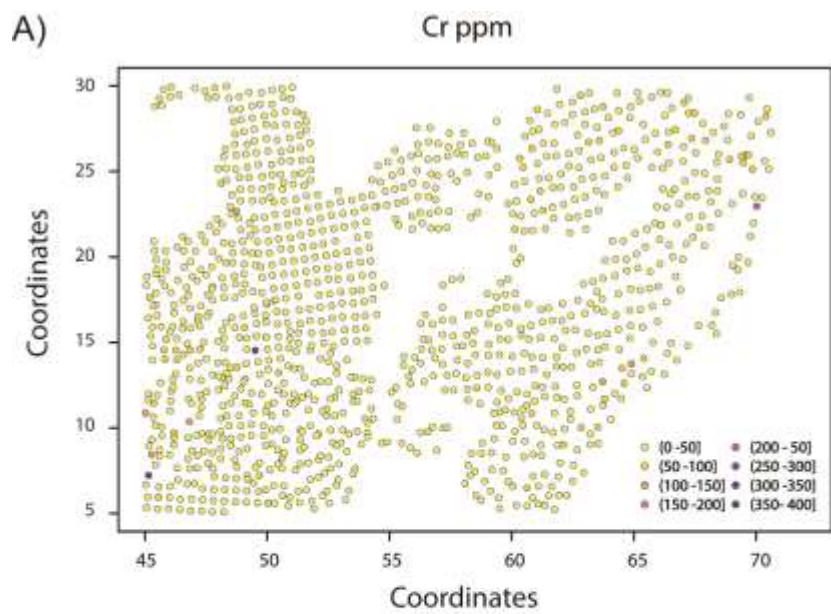
119 elements affected by geological/geochemical processes, which can also be mapped (e.g.
120 principal components). For the first of these tasks, standard practice has included producing
121 a single component map thought to represent the raw or “absolute” input data in the form
122 of dot (or point) maps, but also as interpolated maps of these raw concentrations. Reimann
123 (2005) defines the purpose of such geochemical maps as ‘to display different processes in a
124 map form and to detect local deviations from the dominant process in any one sub-area’.

125 The problems discussed in this section result from the closure property of geochemical
126 compositions, i.e. the unavoidable fact that samples are *aliquots* of the geological bodies we
127 want to investigate and therefore do not really convey information about the element mass
128 distributions (Aitchison, 1986). This has implications for baseline quality mapping, mineral
129 prospecting, land and soil resource assessment or risk evaluation, though these issues are
130 beyond the scope of this contribution. Nevertheless, in the following we show that the
131 traditional meaning of closure effect (linked to closed data with unit, or any other fixed sum
132 constraint of components, and the resulting distortion of the correlation structure) should
133 be considered also in a broader sense. Namely, compositional data are primarily
134 observations that contain quantitatively expressed relative contributions of parts on a
135 whole. From this perspective, the unit sum constraint is just a representation, obtained
136 without altering the source information, conveyed by ratios between the components.

137 Therefore, even with a variable sum of geochemical concentrations, resulting, e.g., from
138 designed omitting of some components, one should be aware that the relative nature of
139 data is still present and needs to be taken into account by proper statistical processing, as
140 exemplified below by CoDA (Pawlowsky-Glahn et al., 2015).

141 **Point maps – the data “as is”**

142 It is often thought that raw one-component maps report “what is there”, that they report a
143 sort of “objective ground truth”. However, Reimann (2005) and others (Reimann and
144 Filzmoser, 2000; Reimann and Garrett, 2005; Reimann et al., 2005; Reimann et al., 2008)
145 highlight that to give sense to that set of spatially dispersed values, space-dependent
146 geochemical processes must be interpreted, and that these are highlighted with a proper
147 representation. Reimann (2005) discusses the advantages and disadvantages of such
148 different representations. Reimann (2005) concludes that, actually, the most important
149 issue becomes the scaling chosen to define the points (colour, size and symbol). Splitting the
150 data into groups (classes) on the basis of order statistics in exploratory data analysis (such as
151 the quartiles of boxplots, or other percentiles), Reimann suggests, may shed light on the
152 spatial structure that reflects at least a number of these processes in a map. Figure 1 shows
153 how impactful this choice of symbol/colour scaling can be with a real dataset. Obviously,
154 equidistant colour scales do not necessarily yield the most easily interpretable maps,
155 neither in raw (or “absolute”) nor in logarithmic values. The proportion of the entire study
156 area actually measured by the dataset is largely exaggerated by the size of the dots. The
157 conclusion is that it may be more appropriate to use quantile-based intervals (following the
158 guidelines on scaling as discussed in Reimann, 2005) to present geochemical data to give an
159 initial assessment of the distribution of elemental concentrations. Readers are referred to
160 Reimann (2005) for further details on methods used to establish intervals in data scaling. It
161 should be pointed out that this choice of scaling does not address the problems resulting
162 from the relative character of elemental concentrations, an aspect that will be explored
163 later.



165 Figure 1: Three representations of Cr ppm concentrations from the GEMAS European soil
166 dataset (Reimann et al. 2014a,b), changing the dependence of the colour scale: (A) using
167 equidistant classes in the raw value scale, (B) using equidistant classes in the log-
168 transformed value scale, and (C) using quantile based (equal probability) classes in the raw
169 value scale. Note that the units are latitude and longitude coordinates but the orientation of
170 the image has been changed to highlight changes that are dependent on the colour scale
171 used.

172

173 **Geospatial continuity – the interpolated map**

174 It is becoming increasingly common for geochemical atlases (online and published printed
175 versions; Reimann et al., 2003, 2014a,b; Young and Donald, 2013) and research publications
176 to produce smoothed (colour) surface maps because most end-users have come to expect a
177 “nice looking” output map. Essential to the generation of an interpolated map, is the
178 demonstration of a spatially coherent pattern through the generation of semi-variograms
179 and other available geostatistical tools. The issues relating to spatial interpolation are well
180 documented elsewhere (e.g. Li and Heap, 2011) and are distinct from the question of the
181 compositional justification for single component maps, which is the focus here.

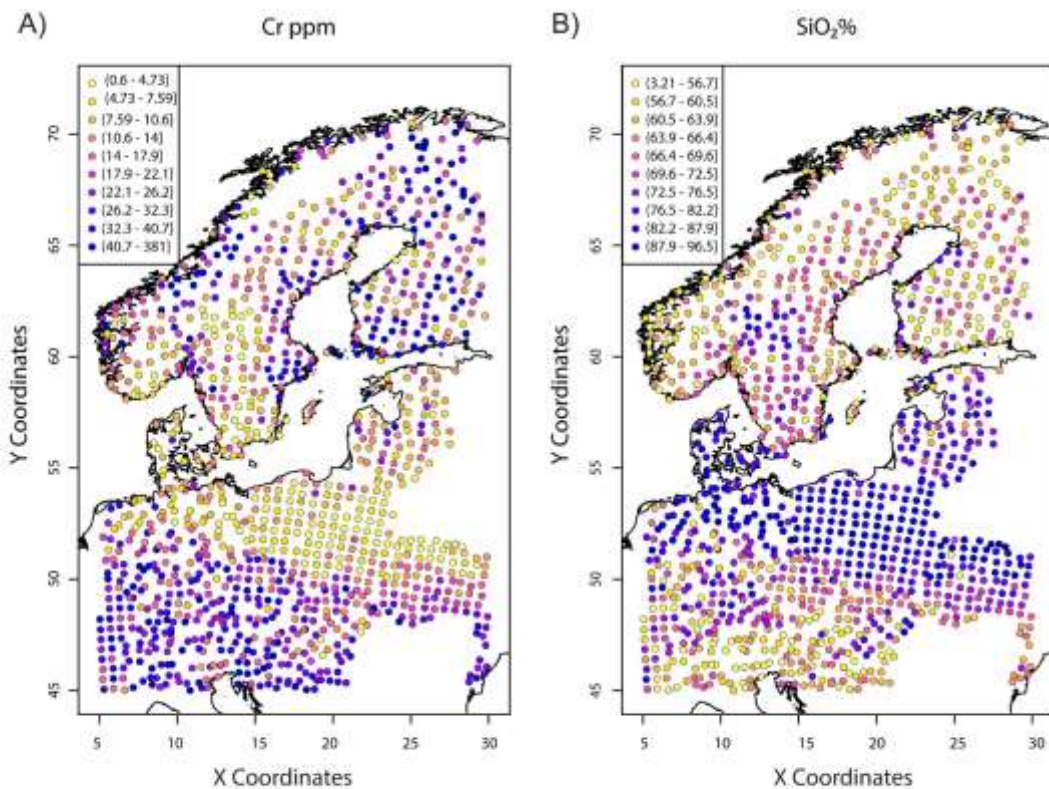
182 Nevertheless, it should be noted that when using raw ‘point data’ the interpolation method
183 should be chosen to ensure several aspects relevant to compositional data, in as much as
184 the interpolated values should (1) remain in the real positive space, and (2) not sum up to
185 more than 100%. Thus interpolated maps should not be obtained separately component by
186 component, as this cannot guarantee adherence to CoDA constraints (as outlined earlier).

187 Joint interpolation methods able to cope with these constraints should be preferred (for
188 example Tolosana-Delgado and van den Boogaart, 2013).

189

190 **Do compositions actually represent absolute or relative abundances?**

191 The central question of this paper is whether a single component map can support an
192 accurate interpretation on its own. As mentioned before, it is often thought that these
193 geochemical maps do actually represent absolute abundances. However, geochemical data
194 are just a particular case of compositional data, and compositional variables carry only
195 relative information of a part to a whole. This can be seen in the conventional concentration
196 units, such as mg/kg, being essentially a constant multiplied by a dimensionless unit: 1
197 $\text{mg/kg} = 1 \text{ mg}/1,000,000 \text{ mg}$, or 10^{-6} mg/mg , or simply 10^{-6} . At every sampled or
198 interpolated point, a concentration value provides information only on the relative weight
199 of one particular element to the total (Tolosana-Delgado & van den Boogaart, 2013),
200 regardless of whether we know all, some or none of the other parts of the whole. This is
201 illustrated in Figure 2 where a lower chromium (Cr) east-west zone is clearly seen in the
202 quantile scaling plot (Fig. 2a). However a corresponding increase in silica oxide (SiO_2) in the
203 same area, (as shown in Fig.2b) shows that any sensible interpretation must consider some
204 sort of possible dilution by SiO_2 that may be related to a change in geology or soil type for
205 example . This is a consequence of the constant sum (closure) effect (Chayes, 1960): if one
206 component increases (due to the addition of a pollutant or a diluent) the mass of all other
207 components remains unaffected, and their apparent proportions (concentrations) must
208 decrease. One solution to this problem is to use element ratios, rather than raw
209 concentrations, because ratios are free of any such effect.



210

211 Figure 2: Illustrative maps from the GEMAS European soil dataset (Reimann et al. 2014a,b):

212 A) (left) map of Cr ppm, B) (right) map of SiO₂ %, showing that the low Cr ppm central band

213 is due to dilution with SiO₂ %. The units are latitude and longitude coordinates.

214

215 Despite this clear limitation, individual raw geochemical maps are deemed fundamental to

216 several applications of geochemical maps, environmental assessments for instance. Thus

217 potential toxicity to humans is typically given as a threshold or a range above which some

218 kind of action is required by law. For example, the Canadian Council for Ministers of the

219 Environment (CCME, 2014) states that arsenic (As) is considered to be potentially hazardous

220 to humans if it exceeds 12-50 mg/kg in soils. The range in tolerance is based on whether the

221 soils are in residential or industrial sites. In Europe a large variety of soil guideline values for
222 agricultural and grazing land soils exists (see Appendix A in Reimann et al., 2014b).

223 Another way of looking at the relative or absolute scale is to consider a simple process
224 affecting a material. Let us imagine a fresh rock with 20% potassium oxide (K_2O) and its in-
225 situ weathered product having 10% of this oxide: what is relevant, that it lost 10% of K_2O or
226 half of it? What would be the most reasonable percentage of K_2O in the weathered material
227 if that original concentration would have been 12%? The correct answer is 6%, and not 2%.
228 These considerations are nevertheless often clearer for trace elements: if a sample has 20
229 ppm of element A and another sample, genetically derived from the first, has 10 ppm of it,
230 half of element A has been removed in the process of forming the second sample. The
231 relative scale effect applies mostly for components with lower concentrations, e.g., for trace
232 elements.

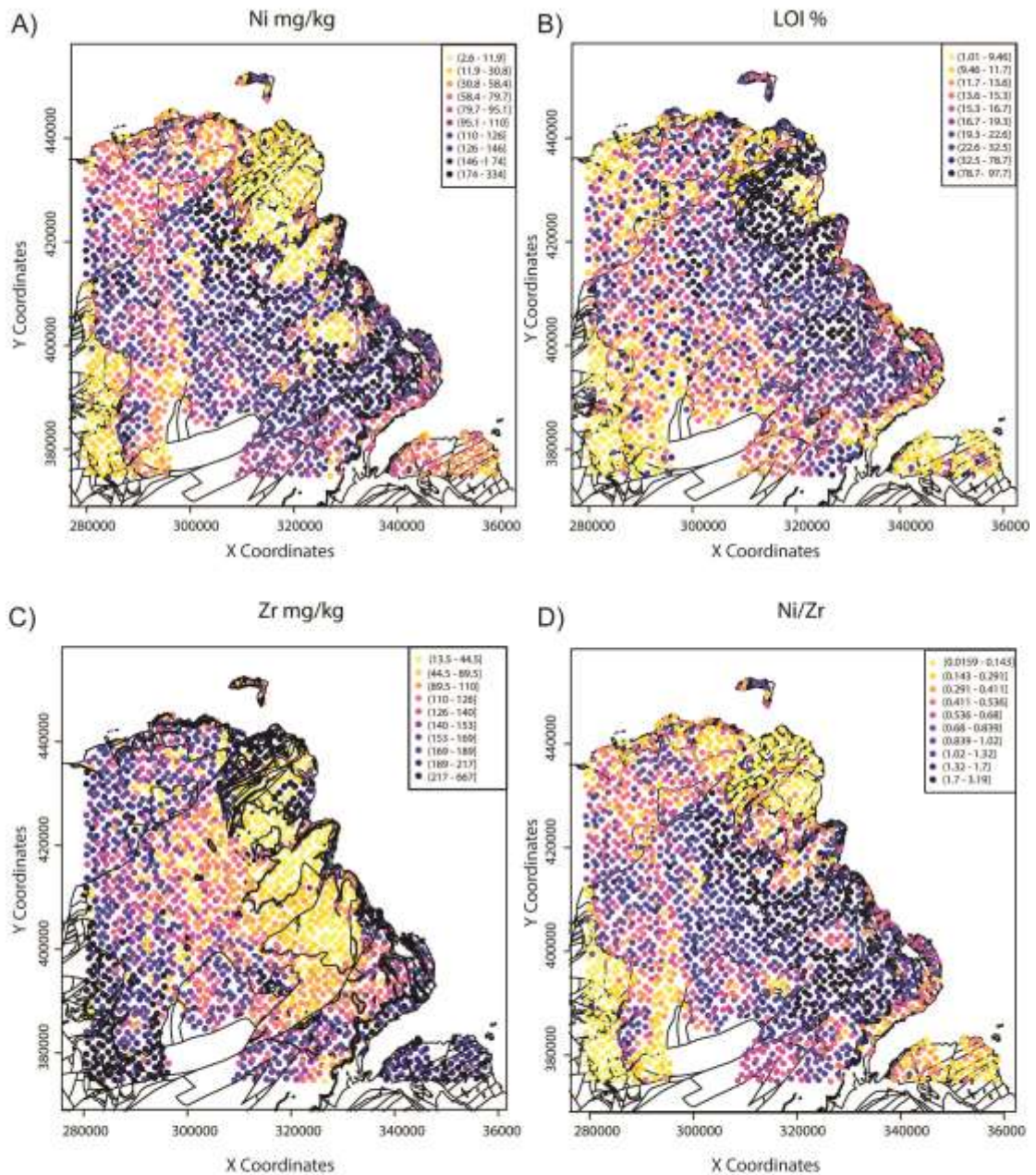
233

234 **Compositional data are always multivariate**

235 In the preceding example where the distribution of Cr was mapped (Figs. 1 and 2) it became
236 evident that any interpretation of the variation in Cr should include an awareness of the
237 dilution effect of other elements and should be done on the basis of ratios such as Cr/SiO_2 .

238 The component increase considered above and its effect on all other components may
239 appear in many other one-component geochemical maps as (high or low) outlier(s) or
240 anomalous value(s). The issue is that geochemical components form a multivariate system
241 that should be analysed as a whole, not component-wise. The following examples from the
242 regional Tellus Survey (Young and Donald, 2013) highlight this aspect. Figure 3 shows the
243 north east region of Northern Ireland where Dalradian metasediments are covered with

244 Palaeogene flood basalts. The basalt bedrock shows high nickel (Ni) concentrations (Fig. 3a)
245 with a low Ni area in the uppermost north east corner related to metasedimentary bedrock.
246 However, lower Ni values are also found in areas where the basaltic bedrock is covered by
247 peat rather than conventional soils (Fig. 3b). A similar effect is seen for zirconium (Zr),
248 though it is enriched in the metasedimentary rocks (Fig. 3c). In areas of consistent bedrock,
249 these low areas of both Ni and Zr are dilution signals formed by high volatiles (represented
250 by Loss On Ignition, LOI %) in blanket peat areas. The variation in geochemical maps can
251 therefore be influenced by dilution, enrichment or depletion in other elements, rather than
252 just by the processes specific to the element mapped. We will show later that maps based
253 on (log)-ratios can avoid this weakness.



254

255 Figure 3: Four map representations using the Tellus soil survey (Young and Donald 2013)
 256 showing the north east region of Northern Ireland: A) Ni mg/kg, B) Loss on Ignition (LOI %),
 257 C) Zr mg/kg and D) Ni/Zr. The units are coordinates in the Irish Transverse Mercator Grid
 258 (EPSG:29903).

259

260

261 **Generalities**

262 A common way to treat the compositional nature of geochemical data has been the
263 application of log-ratio techniques (Aitchison, 1986; Filzmoser and Hron, 2011; Pawlowsky-
264 Glahn and Buccianti, 2011; Reimann et al., 2012; van den Boogaart and Tolosana-Delgado,
265 2013), due to the fact that the ratios between the components are unaffected by constant
266 sum closure effects related to the relative nature of data (Egozcue and Pawlowsky-Glahn,
267 2011).

268 It is important to note that inherent compositional data properties apply to major, minor
269 and trace elements, so there is no reason to distinguish them methodologically, although
270 the effects of closure and other compositional properties (such as the relative scale) depend
271 on abundance and the number of analysed components. In particular, the closure effect can
272 be expected to be higher with a lower number of components and for the major
273 components, while the relative scale will apply mostly in the case of trace elements
274 (Pawlowsky-Glahn and Buccianti, 2011; Egozcue and Pawlowsky-Glahn, 2011; van den
275 Boogaart and Tolosana-Delgado, 2013). Indeed, the closure effect results from distortion of
276 the covariance structure of closed data (Aitchison, 1986), where the sum of covariances of
277 an element to the other components is equal to the negative value of its variance;
278 accordingly, for the extreme case of two-part (closed) compositions the correlation would
279 always equal -1.

280 Advancing towards the conclusions of this paper, the compositional alternative to single
281 component maps will be maps of certain well-chosen log-ratios of two or several

282 components. The next subsections explain the basics of these transformations, followed by
 283 the geochemical and statistical rationale behind each proposed strategy.

284

285 **The log-ratio transformations**

286 Standard logarithmic transformations are not sufficient because it is the ratio to another
 287 element that deals with the closure restraint. The use of one to one ratios does eliminate
 288 the initial closure effect but ratio correlations are known to be subject to restrictions or
 289 “spurious” correlations. Several families of log-ratio transformations exist in the literature.
 290 Aitchison (1986) introduced the pairwise log-ratio transformation (pwlr), the additive log-
 291 ratio transformation (alr) and the centred log-ratio transformation (clr), along with other
 292 transformations which are not relevant for geochemical purposes. Later, Egozcue et al.
 293 (2003) proposed the isometric log-ratio (ilr) transformation. None is inherently better than
 294 the other, each has advantages and disadvantages. The following explanations assume that
 295 the composition $\mathbf{x} = [x_1, x_2, \dots, x_D]$ involves D elements (with boldface indicating a vector
 296 and the square brackets enclosing the concentrations x of elements $1, 2, \dots, D$).

297 The pairwise log-ratio transformation (pwlr) takes *all* possible pairs of elements and
 298 computes their log-ratios.

$$\text{pwlr}(\mathbf{x}) = \begin{bmatrix} 0 & \ln \frac{x_1}{x_2} & \dots & \ln \frac{x_1}{x_D} \\ \ln \frac{x_2}{x_1} & 0 & & \ln \frac{x_2}{x_D} \\ \vdots & & \ddots & \vdots \\ \ln \frac{x_D}{x_1} & \ln \frac{x_D}{x_2} & \dots & 0 \end{bmatrix} = \left[\ln \frac{x_i}{x_j} \right] = [\xi_{ij}]. \quad (1)$$

299 The pwlr contains all the elemental information in compositional data. The treatment of
 300 below detection values and zeros in compositional data is beyond the scope of this paper

301 but readers are directed to Pawlowsky-Glahn and Buccianti (2011), van den Boogaart and
 302 Tolosana-Delgado (2013) and Filzmoser and Hron (2011) for further information on this
 303 subject. This baseline can be used to construct more sophisticated log-ratios, as will be
 304 shown later. The pwlr has the advantage that each transformed coefficient is quite simple
 305 and typically interpretable on its own. Its main inconvenience is the large number of
 306 resulting coefficients, and their high redundancy. For instance even after keeping only one
 307 ratio of each pair of inverse ratios $\ln(x_i/x_j)$ and $\ln(x_j/x_i)$, for $D = 40$ elements, we still need
 308 to consider $D \cdot (D - 1)/2 = 780$ pairwise log-ratios. With regard to redundancy, it can be
 309 seen that any row or column of that matrix can be obtained from additions and subtractions
 310 from the elements of any other row or column, e.g. $\xi_{ij} = \xi_{ik} + \xi_{kj}$ for any three
 311 components i, j, k . Thus, for certain tasks it will be sufficient to take one of the rows or
 312 columns of the pwlr.

313 The pwlr allows us also to introduce an alternative to correlation to be used as a criterion of
 314 statistical association, given the spurious character of raw Pearson correlation coefficients in
 315 a compositional context. This is called the variation matrix, i.e. the variance of each of the
 316 entries of the pwlr matrix: $t_{ij} = \text{Var}[\xi_{ij}]$. The variation is small for pairs of components
 317 which are highly proportional, and it increases as the two components are less and less
 318 dependent on each other.

319 The additive log-ratio transformation (alr) takes just one of the rows or columns of the pwlr
 320 (and removes the constant zero), typically the last column:

$$321 \quad \text{alr}(\mathbf{x}) = \left[\ln \frac{x_1}{x_D} \quad \ln \frac{x_2}{x_D} \quad \dots \quad \ln \frac{x_{D-1}}{x_D} \right] = [\xi_{iD}],$$

322 here arranged in a row to save space. Its advantages are the low number of log-ratios
 323 capturing all the information, and the simplicity of its interpretation. Its disadvantage is the
 324 fact that some results might depend on the chosen common denominator x_D , and therefore
 325 this must be chosen carefully. The choice usually depends on geochemical context (e.g.,
 326 Thomas and Aitchison, 2005), but subjectivity cannot be completely avoided. Another
 327 consideration is that the denominator should be chosen with a minimum number of, and
 328 preferably no, censored values. This ensures that the log-ratio represents a realistic range of
 329 values without interference of instrumentation or limits in the analytical protocols. The alr
 330 transformation has as well certain geometric inconveniences (Egozcue et al., 2003; Egozcue
 331 and Pawlowsky-Glahn, 2006) not relevant for this contribution.

332 Another commonly used transformation is the centred log-ratio (clr) transformation. Instead
 333 of choosing a common denominator, this log-ratio transformation represents each element
 334 component as a ratio to a central value:

$$335 \quad \text{clr}(\mathbf{x}) = \left[\ln \frac{x_1}{g(\mathbf{x})} \quad \ln \frac{x_2}{g(\mathbf{x})} \quad \dots \quad \ln \frac{x_D}{g(\mathbf{x})} \right],$$

336 corresponding to the geometric mean of *all considered* components,

$$337 \quad g(\mathbf{x}) = \sqrt[D]{\prod_{i=1}^D x_i} = \exp\left(\frac{1}{D} \sum_{i=1}^D \ln x_i\right).$$

338 The somewhat artificial role of the geometric mean becomes clear if we consider that each
 339 clr variable represents nothing more than the average of all pairwise log-ratios with one of
 340 the elements, i.e. the averaged sum of a column (or a row) of the pwlr matrix,

$$341 \quad \ln \frac{x_i}{g(\mathbf{x})} = \frac{1}{D} (\xi_{i1} + \xi_{i2} + \dots + \xi_{iD}).$$

342 Note that this averaged sum of all pairwise log-ratios contains the totality of the variability
343 of the elemental composition with respect to the set of components that was considered.
344 Its main advantage is that each transformed score can be related to one component, which
345 might have some interpretability. However, the clr scores always depend on the geometric
346 mean in the denominator, so they cannot be simply interpreted as the original element of
347 interest independently of the other elements. Thus, just as for the individual components,
348 any single component of the clr transform will depend on the notion of the “total”
349 (meaning, the set of all components considered) and can be influenced by dilution,
350 enrichment and depletion in these components. This is the greatest disadvantage of the clr
351 transform. It should be noted that the clr has some desirable geometric characteristics not
352 relevant for this contribution (Egozcue et al., 2003; Egozcue and Pawlowsky-Glahn, 2006). It
353 is important to note that the pwlr, alr and clr transforms do not form orthogonal bases. In
354 the case of the alr transform, the resulting ratios may not reveal the same patterns when
355 there is a change of denominator. For the clr transform, scatterplot matrices will have an
356 intrinsic distortion that may or may not reveal patterns that are meaningful.

357

358 **Log-contrasts and the isometric log-ratio transformation**

359 All transformations mentioned before can be seen as vectors of some scores that are
360 computed as *log-contrasts*, i.e. linear combination of components in log-scale:

$$361 \quad \xi_i = \omega_{i1} \ln x_1 + \omega_{i2} \ln x_2 + \cdots + \omega_{iD} \ln x_D,$$

362 the coefficients of which are subject to $\omega_{i1} + \omega_{i2} + \cdots + \omega_{iD} = 0$. Log-contrasts form
363 special one-dimensional “projections” of compositions. Their result is scale invariant, i.e. it
364 does not depend on the particular units chosen for the components (mg/kg, ppm,

365 proportions, etc.), and have many other desirable stability properties required by CoDA
 366 (Aitchison, 1986). Although log-contrasts seem to be quite complex mathematical objects,
 367 they have a clear geochemical interpretation in connection with chemical equilibrium. If we
 368 assume a reaction $aA + bB \rightleftharpoons cC + dD$, then the logarithm of the equilibrium constant is
 369 one of such log-contrasts, $\xi = \omega_A \ln A + \omega_B \ln B + \omega_C \ln C + \omega_D \ln D$, with $\omega_A = \omega_B = -1$
 370 and $\omega_C = \omega_D = +1$, and it is directly proportional to the Gibbs Energy of the reaction $\Delta G \propto$
 371 $\ln K_{eq}$. Conversely, each log-contrast might be interpreted as a sort of reaction. That might
 372 not be directly meaningful for solid sample geochemistry datasets, but it still offers a way to
 373 interpret log-contrasts. Note that also the components of the above introduced alr and clr
 374 transformations can be seen as log-contrasts. Accordingly, for i -th alr variable ($i =$
 375 $1, \dots, D - 1$) we get

$$376 \quad \ln \frac{x_i}{x_D} = \ln x_i - \ln x_D,$$

377 while i -th clr variable ($i = 1, \dots, D$) can be expressed as

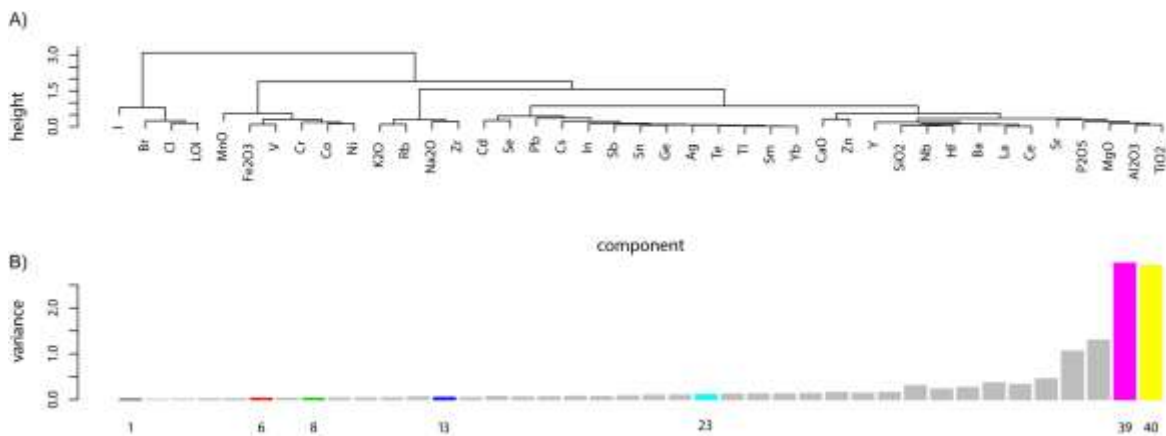
$$378 \quad \ln \frac{x_i}{g(\mathbf{x})} = \ln x_i - \frac{1}{D} (\ln x_1 + \dots + \ln x_D).$$

379 The last log-ratio transformation mentioned is the (family of) ilr transformations (Egozcue et
 380 al., 2003). These are formed by $D - 1$ log-contrasts which are computed using vectors of
 381 coefficients $\boldsymbol{\omega}_i = [\omega_{i1}, \omega_{i2}, \dots, \omega_{iD}]$ which are orthogonal to each other. From all possible
 382 ways of obtaining such orthogonal log-contrasts, the relevant ones for this contribution are:
 383 principal components of the clr-transformed data (Aitchison, 1982; 1986), and balances
 384 (Egozcue et al 2003; Egozcue and Pawlowsky-Glahn, 2005). Balances are simply (normalised)
 385 log-ratios of the geometric means of two groups of elements. One can obtain an ilr based on
 386 balances by choosing a binary hierarchy of association of elements (or a binary partition),

387 such as the one represented in Figure 4. For instance the four log-ratios necessary to
 388 describe the behaviour of the subcomposition (Fe₂O₃, V, Cr, Co, Ni) in this figure are:

389
$$\xi_1 = \frac{1}{\sqrt{2}} \ln \frac{\text{Fe}_2\text{O}_3}{V}, \quad \xi_2 = \frac{1}{\sqrt{2}} \ln \frac{\text{Co}}{\text{Ni}}, \quad \xi_3 = \frac{\sqrt{2}}{\sqrt{3}} \ln \frac{\text{Cr}}{\sqrt{\text{Co} \cdot \text{Ni}}},$$

390
$$\xi_4 = \frac{\sqrt{6}}{\sqrt{5}} \ln \frac{\sqrt[2]{\text{Fe}_2\text{O}_3 \cdot V}}{\sqrt[3]{\text{Co} \cdot \text{Ni} \cdot \text{Cr}}}$$



391
 392 Figure 4: A) An ilr balance dendrogram and B) associated component variance obtained by
 393 choosing a binary hierarchy of association of elements (or a binary partition). The procedure
 394 is detailed in van den Boogaart and Tolosana-Delgado 2013).

395
 396 In general applications of CoDA, the constants of proportionality are necessary to be able to
 397 compare the balances between them. However, in the case of mapping and given that we
 398 are going to use equal-probability intervals (Reimann, 2005), these constants are
 399 superfluous. Interested readers are referred to Pawlowsky-Glahn and Egozcue (2005) and
 400 Pawlowsky-Glahn and Buccianti (2011).

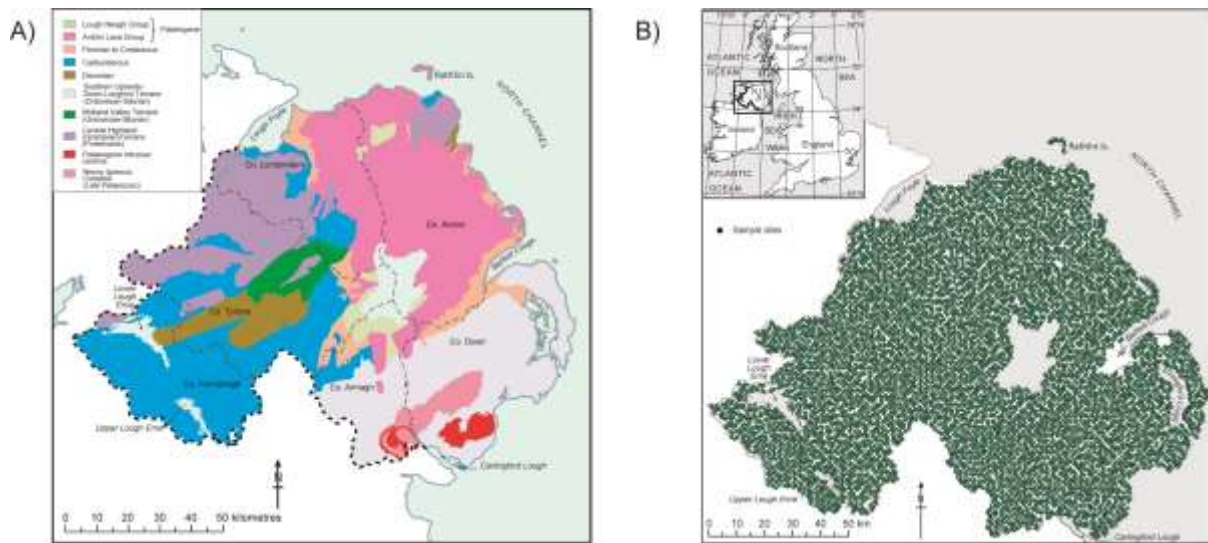
401

402 PROPOSED COMPOSITIONAL ALTERNATIVES

403 Univariate maps are actually produced in the hope that they might have a direct
404 interpretation with respect to specific aspects or processes occurring in the studied region.
405 However, we have argued that this may not be the case, because of the inherent
406 multivariate nature of compositional data and the spurious variations induced by the
407 closure: any process influencing a component will modify all other components (and their
408 maps). As we cannot display or read maps displaying high-dimensional multivariate
409 information, we are left with the question of how to choose interpretable univariate (many
410 single element maps) or low variate (oligovariate) maps (fewer maps displaying two or more
411 variables). All these maps will be representations of certain log-contrasts. The challenge in
412 finding good maps is twofold: there are many more log-contrasts than there are elements in
413 the dataset, and the most appropriate log-contrasts to use will always depend on the aim of
414 the analysis undertaken (Pawlowsky-Glahn and Buccianti, 2011). Using single elements or
415 log-contrasts inherently defines an approach using sub-compositions for gaining insight into
416 geochemical processes. This may work for recognizing obvious processes, but not for
417 processes that are under-sampled or “swamped” by dominant processes.

418 Three approaches for selecting good ratios for various aims are discussed and illustrated
419 with a case study example. These provide alternative approaches to presenting a single
420 elemental component map that honour the compositional nature of the geochemical data
421 and offer interpretable mapped outputs. The feature(s) to be mapped can be based on
422 either (1) a geochemical understanding of the processes under consideration (knowledge-
423 driven log-ratios based on geochemistry), (2) established mathematical projections (e.g.
424 $pwlr$ or clr), or (3) based on the data using statistically determined projections that best
425 capture certain effects, patterns and elemental associations (data-driven projections). These
426 are illustrated using the regional soil sediment geochemical survey covering Northern

427 Ireland, generated as part of the Tellus Survey (Young and Donald, 2013). The soil sediment
428 dataset used comprises 6862 observations of 19 geochemical variables (Fig. 5).



429
430 Figure 5: Case study A) Simplified geology for Northern Ireland (adapted after Mitchell,
431 2004) and; B) Sampling locations for the Tellus Geochemical Survey (GSNI).

432

433 1. Knowledge-driven (log)-ratios

434 In relation to the problem of Ni and Zr dilution by LOI (shown previously in Fig. 3), geological
435 knowledge would suggest that it may be more useful to analyse a simple log-ratio, or a
436 certain balance related to a reasonable equilibrium. Indeed, in this example using the (log-)
437 ratio Ni/Zr reduces the masking effect of peat cover (LOI dilution effect) and illustrates a
438 greater delimitation of the Palaeogene basalt-Dalradian metasediment boundary than any
439 of the three single component maps (Fig. 3d). However, although being of possibly highest
440 interpretational potential, ratio or log-ratio maps present one challenge: we are not used to
441 reading them. This can be resolved by practice and it is advocated that the expected effect
442 of less methodological artefacts promises this investment to be worth the time.

443 Of course, not everything relevant is always expressible in terms of (log-)ratios. In mining
444 there is a need to predict excavation and processing costs per mass recovered, hence the
445 ratio of value element versus mass to be mobilized is relevant, i.e. a quantity proportional to
446 absolute concentration of the value element. In environmental studies, it is important to
447 know how much mass of the toxic elements is present, i.e. proportional to its absolute
448 concentration again. Thus the mass of the element per bulk mass can be a useful proxy.
449 However, in both cases (and in many more examples), other quantities might highlight other
450 important aspects of the problem, quantities which are often ratios or log-ratios. Mining
451 requires the ratio of gangue elements against ore elements, because the product's quality
452 and value depend on them. In environmental issues, well-chosen log-ratios may incorporate
453 bioavailability effects (Barsby et al., 2012) and interactions with other components such as
454 pH or HCO_3^- , etc. Environmental or health hazards typically depend both on absolute values
455 and on ratios, particularly considering bioavailability fractions. Chemical reactions like those
456 mentioned above also play an important role in mobilizing or trapping these elements. For
457 soils and sediments that have been weathered or mixed/diluted with other material, no
458 element shows the same percentages as in the background or source rock, but the ratios of
459 elements unaffected by the mixing (or by pollution or weathering) are preserved. In
460 weathering, especially of magmatic plagioclase-rich rocks, all existing measures of alteration
461 have been defined as (log)ratios, perhaps because quantifying absolute losses of mass
462 would be nonsensical (the removed mass was washed out of the rock-weathered product
463 system, we do not know the starting concentrations of elements in the source rock, the
464 degree of dilution generated from more or less mixing material with the weathered rock is
465 unknown, etc.). Typical practice tends to be to use the ratio of mobile to immobile
466 elements. In studies involving mineral reactions (igneous, metamorphic, diagenesis), Pearce

467 Element Ratios (PER) have been successfully used to model mineral paragenesis (Pearce,
468 1968; Madeisky and Stanley, 2010). In the molar ratio scatterplots used in PER, individual
469 end-member minerals are represented as points and linear trends are interpretable in terms
470 of mineral mixtures. This has been shown to be useful for instance to interpret alteration
471 processes (e.g. Montreuil et al., 2013). Such PER can thus guide the choice of appropriate
472 knowledge-driven log-ratios to map.

473 **2. The pairwise log-ratio map**

474 The second option is to rely on one of the standard log-ratio transformations, for instance
475 the pairwise log-ratio (pwlr). The challenge in this case is the number of pairs that are
476 possible to build: for D elements there are $D \cdot (D - 1)/2$ unique maps. For example the
477 aqua regia analyses from the GEMAS survey result in ~1600 ratio maps, and the XRF
478 compositions add another ~400 ratio maps. It is impossible to present these within a
479 classical atlas publication. Although this is becoming less of a problem with interactive
480 online or electronic atlases, it is not humanly feasible to view and integrate all (or even just
481 many) of these different perspectives. Therefore, either knowledge-driven or data-driven
482 criteria will be required to choose which of this vast number of possibilities are most
483 pertinent for a specific task or problem. While it is acknowledged that in both the
484 knowledge-driven and pairwise approaches, any resultant maps may not explicitly show the
485 impact of other variables relevant for the process being investigated, these approaches do
486 avoid the potential for artefacts due to the constant sum constraint (closure).

487

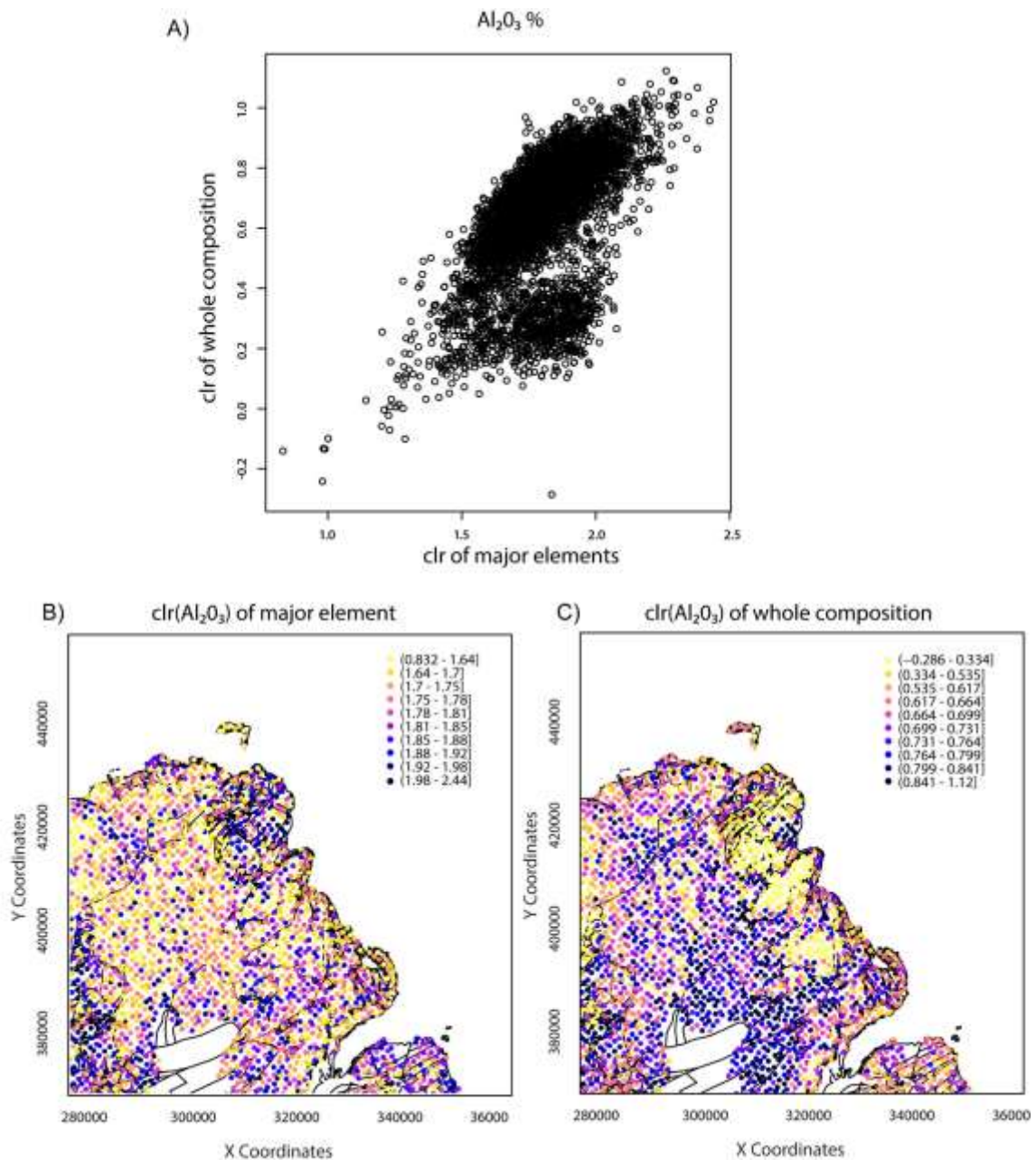
488 **3. The centred log-ratio component map**

489 The usage of the clr transformation may be grounded in one of the following options: (1) to
490 highlight the role of one component, (2) to see the patterns of that element relative to the
491 average behaviour of other elements at hand in the compositional dataset (using their
492 geometric mean), and (3) to provide the first visualization towards a deeper understanding
493 of the underlying geochemical processes.

494 Mapping single clr scores seems to present a first-step simple solution for an unsupervised,
495 blind application. However, it must be noted that clr variables have severe limitations. First,
496 they cannot be interpreted the same way as the variables in the original concentration
497 units, and there is no straightforward back-transformation. A further problem of the clr is its
498 sensitivity to replacement values of censored or missing values, because the more
499 numerous trace elements (with low concentrations and more common censorship)
500 influence the geometric mean more than the fewer major elements (with high
501 concentrations and rarely censored). Therefore, if we impute too low values, we shift the
502 geometric mean to lower values. One solution could be to use a robust central value (e.g.
503 the median of all the components, a trimmed geometric mean, etc.), although this destroys
504 many of the nice properties of the clr and may cause problems for the interpretation of any
505 resulting map. More importantly, as has been mentioned, the clr depends on the set of
506 components chosen, i.e. any component cannot be analysed separately without
507 consideration of the others. This might make interpretation of clr maps difficult, especially
508 when the other components are driven by their own geochemical processes.

509 The above features are illustrated using the Tellus dataset. Figure 6a demonstrates the
510 effect of employing the whole composition or the subcomposition of major oxides,
511 respectively, for clr scores of these elements. In addition to a shifting of the distribution, a

512 result of excluding the trace elements, the positive linear correlation between values of
513 both scores is substantially disturbed for some components (e.g. in particular for Al_2O_3 ,
514 MgO , SiO_2). This is also reflected in the corresponding univariate maps (as in the case of clr
515 Al_2O_3 , see Figure 6b and c). This clearly demonstrates that the selection of the
516 subcomposition is important for the construction of clr scores. Here the use of the
517 subcomposition of major oxides clr scores seems to be reasonable due to their higher
518 geochemical stability in the study area. On the other hand, by employing all elements we
519 can observe some clear regional patterns, that may indicate interesting inherent processes
520 in the data that are not so well reflected from the major oxides subcompositional clr.



521

522 Figure 6: A) The effect of employing the clr of the whole composition and the clr of the
 523 subcomposition of the major oxides is illustrated for Al₂O₃. Note the substantial
 524 disturbance in the positive linear correlation. The corresponding univariate maps
 525 demonstrating the effect of employing B) the clr (Al₂O₃) of the subcomposition of major
 526 oxides or C) the clr (Al₂O₃) of the whole composition is shown. The units are coordinates in
 527 the Irish Transverse Mercator Grid (EPSG:29903).

528

529 In the authors' opinion, using the geometric mean of a geologically meaningful subset of
530 components, selected on the basis of a geological understanding, may offer a possible way
531 to form a compromise between the two situations that we have demonstrated (Fig. 6 b and
532 c). That is using pairwise log-ratio maps (one element to another, or any knowledge-driven
533 log-ratio) and using the clr scores of the full composition including the problematic
534 components (i.e. those having a high proportion of values that are censored, close to the
535 detection limit, or rounded). The result may offer a reasonable number of single (clr score)
536 or low) component maps. For the Tellus Survey example it would seem a sensible approach
537 to include the major oxides Al_2O_3 , CaO , Fe_2O_3 , K_2O , MgO , MnO , Na_2O , P_2O_5 , SiO_2 and TiO_2 ,
538 with the exception of LOI and SO_3 to avoid problems with the soil water content and with
539 the large number of missing values of SO_3 . Their subcompositional clr-scores could then be
540 mapped. For the remaining elements, comprising the trace elements, LOI and SO_3 , the
541 balance of each element against the geometric mean of the major component
542 subcomposition could be mapped.

543 Another use of a subcompositional clr is to restrict attention to a knowledge-driven
544 selection of parts, and constructing and mapping their subcompositional clr scores. If the
545 rationale for the study is set out beforehand, then these clr maps of subcompositions should
546 be interpretable, because the subcomposition was selected and studied for a particular
547 reason. An example of this approach could be to use a subcomposition of potentially
548 harmful elements that are known to adversely affect human health (for example As, Cr, Pb
549 and Cd). In fact, this approach is reminiscent of the first method, knowledge-driven log-

550 ratios, where the interpretability of the resulting variables is assumed since there was a
551 knowledge-driven rationale for the approach.

552

553 **4. Exploring patterns and elemental associations (data-driven projections)**

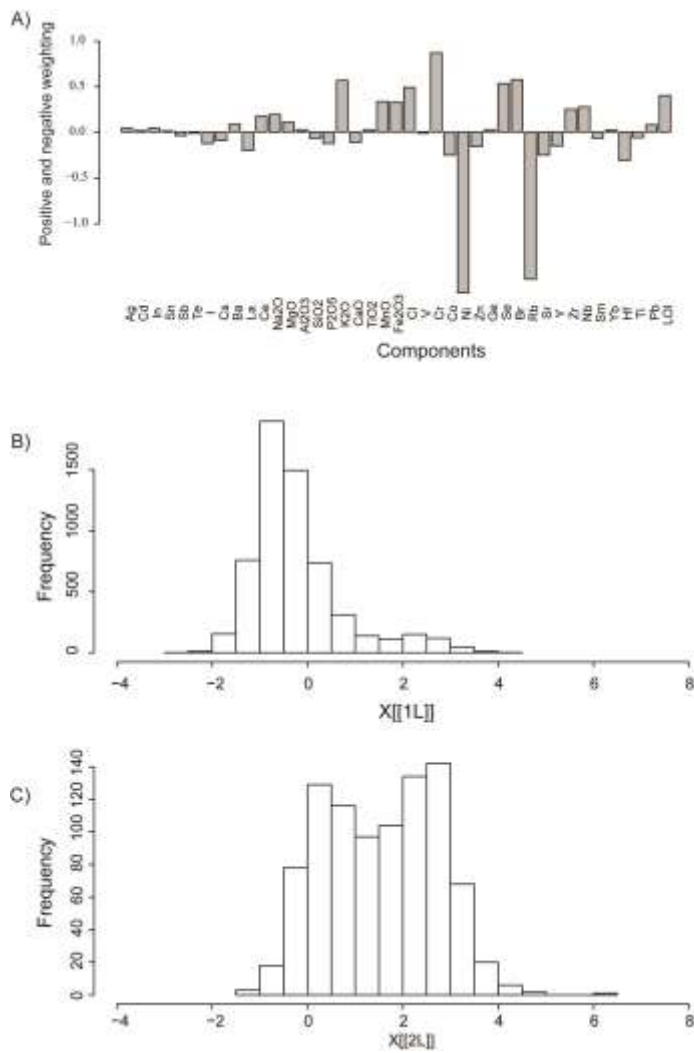
554 Multi-element geochemical datasets are often analysed to explore patterns, links and
555 associations between elements as part of the 'Discovery Process' (Grunsky et al., 2014).

556 Multivariate methods available include principal component analysis (PCA), regression-
557 related techniques (regression analysis, analysis of variance, total and partial least squares
558 regression or canonical correlation) and grouping techniques (cluster analysis, discriminant
559 analysis), often used to explore these patterns or links. From the point of view of mapping,
560 most of these methods can actually be seen as a way to choose one (or a few) log-contrasts
561 of potential interest. There is thus a range of techniques available and the issue is to choose
562 the best one. Two families of approaches can be considered: supervised and unsupervised
563 methods. Again, these will be exemplified using the Tellus case study.

564 *Supervised Approaches*

565 In the supervised approach the goal is to find which log-contrast better relates to known
566 secondary information which is also available. For example, we may be interested in
567 identifying the log-contrast which best distinguishes between two types of geochemical
568 situations such as two terranes, or areas of peat coverage and peat free areas. In that case
569 we could start from a subset of the data, where this extra information is known (i.e. a
570 *training set*). The best log-ratio can be computed by a linear discriminant analysis derived
571 from any log-ratio transformed composition, although due to numerical issues, $\ln r$ and $\ln R$
572 are the best in this case (readers are referred to van den Boogaart and Tolosana-Delgado,

573 (2013) for further detail). Such a linear discriminant analysis has been conducted on the
574 Tellus soil geochemical dataset on the basis of areas covered by blanket peat (Figure 7). The
575 linear discriminant analysis suggests that many more elements and oxides (Ni, Rb, Hf, Nb,
576 Cr, K₂O, MnO, Fe₂O₃, some with positive weight, some with negative weight; Fig.3a) are
577 useful for blanket peat delineation than solely using LOI %, as is traditionally the case.
578 Overall, the histograms shown in Figure 7b and 7c show that the classification between peat
579 covered and peat free areas, produced by geostatistical Fisher discriminant analysis in this
580 example (Tolosana-Delgado et al. 2015), is reasonably successful. On further investigation it
581 is observed that misclassified points are found mainly at the boundaries between peat-
582 covered areas and peat-free areas and most likely related to point-polygon mismatch
583 between the Tellus data points and peat assigned polygons (discussed in more detail in
584 Tolosana-Delgado et al. 2015). A full discriminant analysis approach could be applied in that
585 results could be validated further using cross-validation, the log-contrast could be derived
586 from a training set and the discriminant power of the analysis could be evaluated.



587

588 Figure 7: Illustrating a supervised method: Linear discriminant analysis using the Tellus soil
 589 geochemical dataset on the basis of blanket peat; A) elemental linear discriminant analysis,
 590 showing positive and negative weightings and histograms showing the classification
 591 between B) peat covered and C) peat free areas, produced by the geostatistical Fisher
 592 discriminant analysis. Readers are referred to Tolosana-Delgado et al. (2015) for further
 593 detail.

594

595 In a similar way, if the covariable(s) of interest is (are) a numeric quantity or a set of them
 596 (magnetics, gravimetrics, Th-channel radiometrics, total gamma-ray counts, proportion of

597 sand, log-ratio of silt to sand, pH, etc.), compositional regression analysis could be used in
598 order to identify the optimal log-contrast (e.g McKinley et al. in review). The composition
599 should be expressed in log-ratios (alr or ilr) and a classical regression applied between the
600 log-ratio scores and the chosen covariable. The explanatory power of the resultant model
601 can be measured by its associated R-square. Thus, beyond the classical use of regression, it
602 is possible to use this criterion to select the optimal pairwise log-ratio that maximizes the R-
603 square of the simple regression model with the covariable of interest. This requires the
604 exhaustive exploration of all possible pairwise log-ratios, though the selection of the “best”
605 is automatically done by the method and not by visual inspection. Note that these
606 techniques, although explained for one covariable or two groups, can be easily generalized
607 towards multiple groups or covariables, providing an extra direction for each added group
608 or covariable is known (van den Boogaart and Tolosana-Delgado, 2013).

609

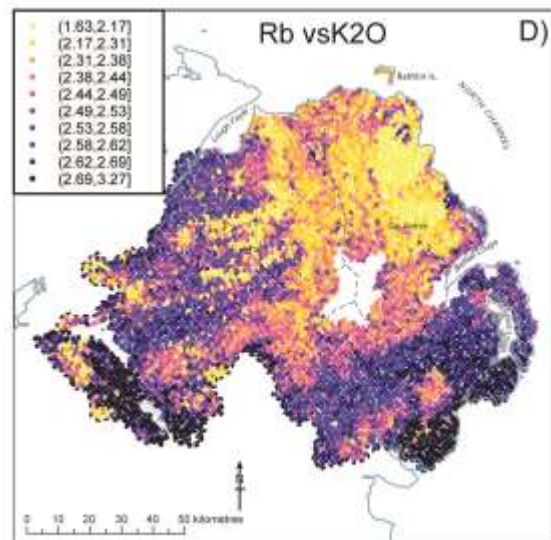
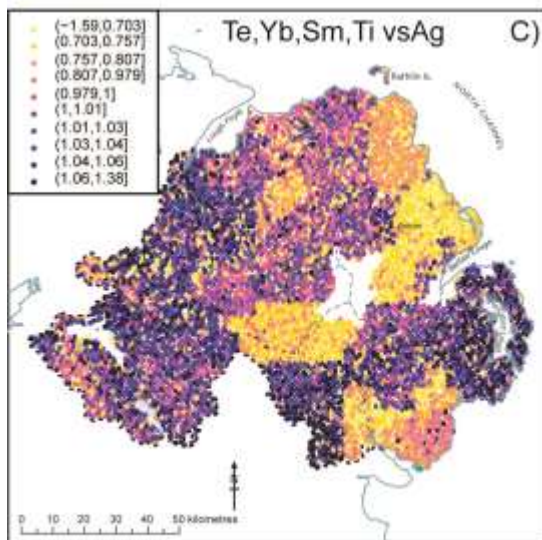
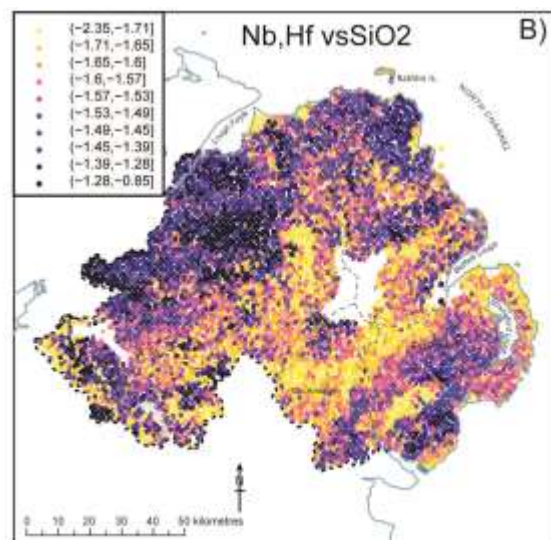
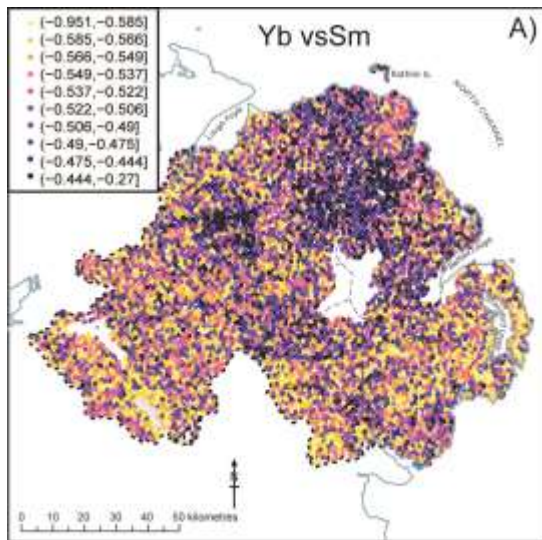
610 *Unsupervised approaches*

611 Even in the absence of additional information, multivariate analysis techniques can help to
612 identify interesting log-contrasts to map. These often show either large variability or very
613 low variability. For example, the variation coefficient t_{ij} of two components provides a
614 measure of their association, which can be used in cluster analysis as a measure of distance
615 between variables. The application of cluster analysis for compositional data (compositional
616 Q-mode clustering) is explained in detail by van den Boogaart and Tolosana-Delgado (2013).
617 Clusters will contain elements behaving proportionally throughout the dataset. Log-ratios
618 between elements of two different clusters should thus be similar to other log-ratios of the
619 elements of the same clusters. Therefore, one of these log-ratios or a balance of one cluster

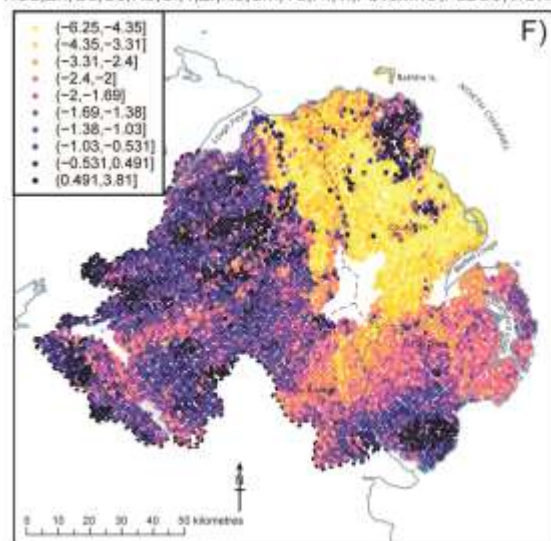
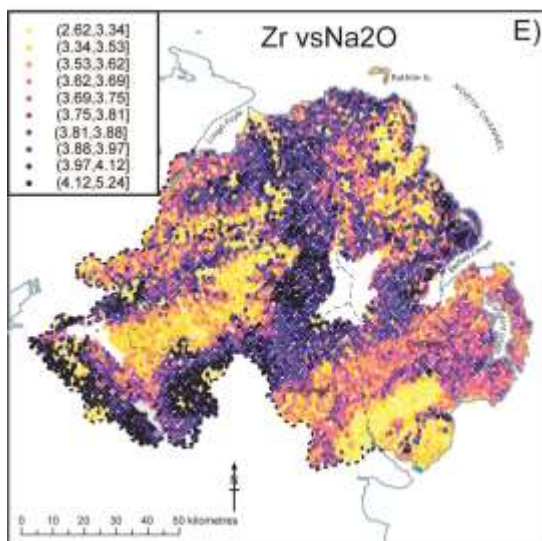
620 against the other might be representative for many log-ratios, and consequently may
621 represent a process influencing many elements in the same way. Balances of elements
622 within the cluster will filter out these large-variability effects and focus on differences
623 between elements behaving similar with respect to major processes. In a sense this is the
624 approach that geochemists instinctively use when applying simple ratios. The advantage of
625 this unsupervised CoDA approach that reflects a classical supervised approach is that these
626 balances should thus be able to highlight processes differentiating between similar
627 elements, i.e. second-order processes that would otherwise be overshadowed by the major
628 processes.

629 In the example shown in Figure 8 using the Tellus regional soil data, a cluster dendrogram is
630 constructed using all those components without missing values (shown in Figure 4).

631 Although it is acknowledged that this introduces a bias, since the treatment of below
632 detection values and zeros in compositional data is beyond the scope of this paper, this
633 simplified approach is used here to exemplify the approach. The hierarchical cluster has
634 been produced with a Ward cluster analysis with the variation matrix as distance between
635 elements, the procedure is detailed in van den Boogaart and Tolosana-Delgado (2013). Six
636 resultant ilr balances have been shown to demonstrate different features: spatial
637 randomness (Fig. 8a; Yb vs Sm); the identification of granitic bedrock areas (Fig. 8b; Nb, Hf
638 vs SiO₂); an artefact in the central region perhaps related to the analytical method of Ag
639 detection (Energy Dispersive Polarised X-Ray Fluorescence, Fig. 8c, Te, Sm, Yb, Tl vs Ag);
640 Geological bedrock units including Palaeogene basalts, a granodiorite complex and
641 Dalradian metasediments (Fig. 8d, Rb vs K₂O); Ordovician- Devonian sandstone units and
642 granitic rocks (Fig. 8e, Zr vs Na₂O); Palaeogene basalt detection and LOI-related peat
643 unsupervised detection (Fig. 8f).



Ag, Cd, In, Sn, Sb, Tr, Cs, Ba, La, Ce, Na₂O, MgO, Al₂O₃, SiO₂, P₂O₅, K₂O, CaO, TiO₂, Zn, Ge, Se, Rb, Sr, Y, Zr, Nb, Sm, Yb, Hf, Ti, Pb vs MnO, Fe₂O₃, V, Cr, Co, Ni



645 Figure 8: The figures illustrate an unsupervised method based on the ilr balance
646 dendrogram shown in Figure 4: Six resultant ilr balances are selected to illustrate different
647 features: A) spatial randomness (Yb vs Sm); B) granitic bedrock areas (Nb,Hf vs SiO₂); C) an
648 artefact related to the analytical method of Ag detection (Te,Sm,Yb,Tl vs Ag); D) Geological
649 bedrock units including Palaeogene basalts, granodiorite complex and Dalradian
650 metasediments (Rb vs K₂O); E) Ordovician- Devonian sandstone units and granitic rocks (Zr
651 vs Na₂O ; F) Palaeogene basalt detection and LOI-related peat detection.

652

653 Another popular approach traditionally used for identifying interesting log-contrasts is
654 principal component analysis of the clr-transformed dataset and its graphical display, the
655 compositional biplot (Aitchison and Greenacre, 2002; Caritat and Grunsky, 2013). Within a
656 compositional approach this allows the identification of directions or log-contrasts of large
657 or small variations based on a variance-covariance matrix. Log-contrasts of strong variation
658 might be useful for identifying major processes strongly influencing many elements at the
659 same time. The directions of small variation are more useful for identifying common
660 equilibrium-like laws and anomalies by highlighting cases showing strong deviations from
661 otherwise stable log-contrasts. These directions of lesser magnitude may also represent
662 random processes in the data. Any of the methods mentioned before can be used in several
663 different ways with the purpose of focussing on specific features and aspects of the dataset.
664 In summary, there are many possible tools and approaches for selecting informative and
665 valid univariate projections of a composition. The key point is to relate them to specific
666 research questions to address pertinent geochemical and geological issues.

667

CONCLUSIONS - BEST PRACTICE AND RECOMMENDATIONS

668 The question that this paper has addressed is whether the single component elemental map
669 is meaningful: can it be interpreted as fact or is it indeed no more than a fictional
670 representation? The essence of this dilemma is that geochemical data are compositional
671 variables and as such carry only relative information. In some cases, the classical single-
672 component maps might carry some relevant information, but most of the time the
673 information-rich relative data are not properly represented. On reflection, this research
674 recommends the following:

6752. An appropriate manner to represent "measured" raw or absolute geochemical values is in
676 the form of a single component point source or 'dot maps' with an acknowledgement about
677 dependence to all other variables. This may be useful for the purpose of geochemical
678 exploration but is critical for environmental studies where stated values fall below/above
679 guidelines. Generally, without considering or knowing the other elements of the
680 composition, we are not able to fully interpret the component of interest. That is to say that
681 involving more components in the interpretation can reveal the reason behind the relative
682 increase or decrease in the component of interest.

6833. A complementary compositional approach is required to ensure that mapped outputs are
684 accurate and robust to interpret the underlying geochemical processes and highlight any
685 anomalies of interest. One suggestion is to look at log-ratio maps, which may include either
686 knowledge-driven log-ratios, pairwise log-ratios, any reasonably aggregated log-ratios or
687 data-driven projections, in addition to single element maps whenever possible. In this case,
688 the authors propose the following chain of representations:

- 689 • in paper format, the maps of clr-transformed major components together with the
690 balance of each trace component to the major components; in electronic interactive
691 formats, this can be complemented with maps of each pwlr score;
- 692 • maps of clr- or ilr-based principal components, those with highest variability and
693 also those with very low variability; it is worth noting that one should be careful if
694 including all elements for principal component analysis as trace ones (resulting in
695 highest variability due to relative scale of compositions) are often burdened by
696 measurement errors;
- 697 • maps of the balances resulting from a compositional cluster analysis based on the
698 variation matrix.

6994. No matter whether absolute or log-ratio maps are used, most commonly it will be necessary
700 to use quantile-based colour and symbol scales to highlight the patterns captured. In log-
701 ratio maps, equal length intervals for colours and symbols are often useful as well. However
702 this is close to useless in absolute maps except to highlight a few high-value outliers.

7035. If a compositional approach is not used, it becomes difficult to guarantee that any identified
704 pattern, trend or anomaly is not an artefact of the constant sum constraint. Nevertheless, it
705 should be noted that if the elements do not sum up to the same constant, the relative scale
706 of compositions still needs to be taken in account.

7076. The evaluation of individual component maps is a first step towards a comprehensive
708 understanding of geochemical processes that are reflected in the material that has been
709 analysed geochemically. It is the initial step in the “Discovery Process” from which
710 geologically meaningful log-contrasts can be derived. There are so many possible supervised
711 and unsupervised data-driven methods, that it is necessary to: (1) clearly state the

712 geological/geochemical question to ask; (2) search for the appropriate statistical method
713 that can answer that question; (3) ensure that the data available are sufficient to answer
714 that question with the selected method; (4) apply the required log-ratio transformations
715 and the chosen statistical method; and (5) learn to interpret the results. This will most often
716 require statisticians, data analysts and geochemists working together.

717

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