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1 **Title:** Online control of reaching and pointing to visual, auditory, and multimodal targets:
2 Effects of target modality and method of determining correction latency

3

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17

18 **Running head:** Multimodal online control

19 **Abstract**

20 Movements aimed towards objects occasionally have to be adjusted when the object
21 moves. These online adjustments can be very rapid, occurring in as little as 100ms. More is
22 known about the latency and neural basis of online control of movements to visual than to
23 auditory target objects. We examined the latency of online corrections in reaching-to-point
24 movements to visual and auditory targets that could change side and/or modality at
25 movement onset. Visual or auditory targets were presented on the left or right sides, and
26 participants were instructed to reach and point to them as quickly and as accurately as
27 possible. On half of the trials, the targets changed side at movement onset, and participants
28 had to correct their movements to point to the new target location as quickly as possible.
29 Given different published approaches to measuring the latency for initiating movement
30 corrections, we examined several different methods systematically. What we describe here
31 as the optimal methods involved fitting a straight-line model to the velocity of the correction
32 movement, rather than using a statistical criterion to determine correction onset. In the
33 multimodal experiment, these model-fitting methods produced significantly lower latencies for
34 correcting movements away from the auditory targets than away from the visual targets. Our
35 results confirm that rapid online correction is possible for auditory targets, but further work is
36 required to determine whether the underlying control system for reaching and pointing
37 movements is the same for auditory and visual targets.

38 **Keywords:** Multisensory, multimodal, space, online control, methods

39 **1 Introduction**

40 When reaching to point towards or grasp an object, it occasionally moves
41 unexpectedly, or we dislodge it with our hand, or our initial movement was inaccurate. We
42 then have to correct our movement 'online' during its execution. Online control may be the
43 default mode of visuo-motor control, rather than using a model-based or predictive form of
44 control (Zhao & Warren, 2015). Online movement corrections can be very rapid. In cats
45 reaching for a food reward, paw movements can be corrected within as little as 60-70ms
46 following changes in target location (Alstermark, Eide, Górska, Lundberg, & Pettersson,
47 1984). In humans, significant changes in reaching movement acceleration have been
48 reported as early as 90ms after the target displacement (Paulignan, MacKenzie, Marteniuk,
49 & Jeannerod, 1991). The online control of movements has been thoroughly investigated for
50 changes in the location, size, and other features of visual targets (Paulignan et al., 1991a;
51 Paulignan, Jeannerod, MacKenzie, & Marteniuk, 1991; Veerman, Brenner, & Smeets, 2008;
52 Wijdenes, Gomi, and Brenner, 2015), but the online control of movements towards auditory
53 targets has only just begun to be studied (Boyer, Babayan, Bevilacqua, Noisternig, Warusfel,
54 Roby-Brami, Hanneton, & Vivaud-Delmon, 2013; see Cameron and López-Molinar, 2015,
55 Cluff, Crevecoeur, & Scott, 2015, for similar points regarding proprioception). The present
56 study investigated the ability of healthy human participants to make online movement
57 corrections to visual, auditory, and multimodal targets. In particular, we compared the
58 latencies of these corrections. By 'multimodal' target, we mean a target that begins either as
59 visual or auditory, then switches modality after movement onset, to become auditory or
60 visual, respectively.

61 The online control of movements to visual targets is thought to be a function of the
62 dorsal visual stream: damage to the superior occipital-parietal cortex impairs the online
63 control of reaching movements (Pisella, Gréa, Tilikete, Vighetto, Desmurget, Rode, Boisson,
64 & Rossetti, 2000), and targets thought to be processed most rapidly by the magnocellular
65 pathway of the dorsal stream are associated with lower latency online control (Veerman et

66 al., 2008). There is little evidence concerning the neural basis of online control of movements
67 towards auditory targets. In macaques, neurons in the parietal and premotor cortices may
68 represent the locations of targets across modalities in a common reference frame, for eye
69 and hand movements (Cohen & Anderson, 2000; Graziano, Reiss, & Gross, 1999). Further,
70 the superior colliculus, which receives inputs from vision, audition, and somatosensation, as
71 well as other brain stem regions has been implicated in the online control of reaching
72 movements in the cat (Alstermark et al., 1984; Pettersson, Lundberg, Alstermark, Isa, &
73 Tantisira, 1997), and in primates (Song, Rafal, & McPeck, 2011; Werner, 1993).

74 Given that these brain areas thought to be involved in the online control of movement
75 are responsive to multiple sensory modalities, we speculated that some aspects of the online
76 control of movements may be multimodal or supramodal in nature, and, further, that rapid
77 online control may even be possible for targets that change modality as well as location.
78 Changes in target modality such as this might occur in nature, for example with a cat chasing
79 a mouse (Alstermark et al., 1984): As the mouse runs behind an object, it is visually occluded
80 from the cat, but auditory cues may still be available to guide pursuit.

81 We asked healthy volunteers to make speeded reaching and pointing movements to
82 visual (Experiment 1) and auditory (Experiment 2) targets, which changed location on 50% of
83 the trials (from left-to-right or right-to-left), and, in the third experiment, orthogonally could
84 also change modality (from auditory-to-visual or visual-to-auditory) after movement onset.
85 We determined the time-point at which the movement trajectory changed in the different
86 conditions. Following reviewers' comments, we systematically investigated two different
87 methods of determining latency (statistical, and extrapolation), for three different levels of
88 analysis (whole group, individual participant, and individual trial), and three different types of
89 velocity (lateral, resultant, and statistical components of velocity) – 18 different combinations.
90 For the statistical methods, 61 different statistical thresholds were assessed. This systematic
91 investigation allowed greater certainty in our conclusions, but also highlighted large
92 differences between different methods of estimating correction latency from velocity data.

93 To summarise, we aimed first to compare different methods of measuring correction
94 latency (see also Wijdenes, Brenner, and Smeets, 2014), second to examine the latency of
95 online corrections for pointing to auditory targets in comparison with visual targets, and third
96 to examine the latencies of movement corrections made to both visual and auditory targets
97 that can change modality and/or position at movement onset.

98

99 **2 Materials and Methods**

100 **2.1 Participants**

101 Thirteen participants (7 male, 6 female; 11 right-handed; aged between 20 and 33
102 years; including two of the authors) took part in the experiments. All of the participants had
103 normal or corrected vision. All participants gave written, informed consent, the experimental
104 procedures were approved by the local ethical review panel at the Hebrew University of
105 Jerusalem, and were in accordance with the Declaration of Helsinki (as of 2008).

106

107 **2.2 Apparatus and materials**

108 The experiments were performed in a darkened sound-attenuated chamber (Eckel C-
109 26, UK). Participants sat in the middle of the chamber on a straight-backed chair with a
110 horizontal board as a forearm rest supporting a small marker for the starting position in the
111 centre of the chamber (Figure 1). An arced metal hoop of 90cm radius supported an array of
112 three loudspeakers (7.5 degrees left, centrally, and 7.5 degrees right of the midline), and
113 three 5mm diameter LEDs (left, centre, and right, attached centrally in front of each
114 loudspeaker).

115 Index fingertip and head position (3 degrees of freedom) and orientation (3 degrees of
116 freedom) were recorded with a Polhemus Patriot (Polhemus, Colchester, VT, USA) magnetic
117 tracking system, sampling at 60Hz. The transmitter was positioned centrally, in front of and
118 below the participant, between their knees. Participants wore plastic goggles which held the
119 head position tracker and a laser pointer, used to assist calibration of head position prior to

120 data collection. Horizontal and vertical electrooculographic (EOG) data were acquired with an
121 Active 2 Biosemi system (Biosemi, Amsterdam, The Netherlands), sampling at 1024Hz, with
122 an online low-pass filter of 256Hz. Four electrodes were used for EOG recording: two
123 electrodes for the horizontal EOG, at the outer canthi of the left and right eyes (HEOGL,
124 HEOGR), and two vertical EOG electrodes below (infraorbital, VEOGI) and above
125 (supraorbital, VEOGS) the right eye. Two channels were recorded from the mastoid
126 processes and another from the tip of the nose, but were not used. Bipolar EOG channels
127 were created offline by subtracting HEOGL from HEOGR and VEOGI from VEOGS. The
128 data were referenced online to a common-mode.

129

130 **2.3 Stimuli**

131 Visual and auditory stimuli were presented by passing the same amplitude- modulated
132 white noise stimulus waveform through the sound card of a PC. A parallel port signal
133 triggered a relay switch box that channelled the stimulus to either a loudspeaker or an LED
134 (5mm, red, ~800 mcd). The stimulus was generated on each trial as follows: A 1250ms white
135 noise signal sampled at 44,100Hz, ranging from -1 to +1 was attenuated by 5% to prevent
136 clipping, and shaped with a trapezoidal envelope providing 10ms rise and fall times. To
137 facilitate the perceptual localization of the auditory stimuli, the stimulus was multiplied by a
138 sinusoidal envelope with a frequency of 60Hz, providing an amplitude-modulation depth of
139 80%. Thus, perceptually, the auditory stimulus fluttered, while the visual stimulus was only
140 just visibly flickering if viewed peripherally. In pilot work, participants could easily discriminate
141 between auditory stimuli as little as 7.5 degrees apart (the closest that the speakers could be
142 put together without touching). The 95% confidence intervals for directional pointing error
143 were approximately 5 degrees in azimuth and elevation. Auditory stimuli were presented at a
144 mean approximately 75dB sound pressure level, but to reduce the possibility that auditory
145 stimulus intensity acted as a cue to distance or direction, stimulus amplitude was varied
146 randomly on each trial with a rectangular distribution between 75% and 100% of the original

147 amplitude. A central visual fixation stimulus (5mm diameter LED), remained on until the hand
148 and head were correctly positioned at the start of each trial. An auditory preparatory cue
149 (1000ms, 960Hz, central loudspeaker) was used to signal the start of each trial and to cue
150 the participants to fixate. Trials ended with the illumination of the central LED, and
151 presentation of a 1000ms 480Hz tone from the central speaker.

152

153 **2.4 Design**

154 Three experiments were performed in a pseudo-randomised fashion. The first
155 contained only visual targets (96 trials in total), the second only auditory targets (96 trials),
156 and the third (the multimodal experiment) contained visual, auditory, visual-then-auditory,
157 and auditory-then-visual targets (192 trials). The three experiments were performed on the
158 same day in a single session lasting 2-3 hours. The multimodal experiment was always run
159 either first (6 participants) or last (6 participants), while the order of the visual only and
160 auditory only experiments was counterbalanced across participants. The first participant did
161 not participate in the multimodal experiment because of technical problems.

162 Within the visual only and auditory only experiments, four conditions were generated by
163 the exhaustive combination of two binary variables: Initial target position (left or right), and
164 final target position (left or right). Thus, in half of the conditions, the target remained on one
165 side, and in the other half, it switched from one side to the other. The multimodal experiment
166 contained 16 conditions, formed from the exhaustive combination of four binary variables:
167 Initial target position (left or right), final target position (left or right), initial target modality
168 (visual or auditory), and final target modality (visual or auditory). Thus, in half of the trials,
169 there was a change in target position, and orthogonally in half of trials there was a change in
170 target modality at movement onset.

171

172 **2.5 Procedure**

173 Participants were briefed and gave informed consent, then the EOG electrodes were
174 attached. The participants were seated in the chamber and the tracking devices were
175 attached. 12-24 practice trials were performed before data collection in each experiment to
176 familiarise the participants with the task. The chamber was kept nearly dark, but due to
177 ambient light and the LEDs, the participants received rudimentary visual feedback of their
178 hand and arm positions during the experiment. To prevent complete dark adaptation, the
179 chamber door was opened every 10-15 minutes, between blocks of trials.

180 The participants were instructed to reach and point, as quickly and as accurately as
181 possible with the index finger of their dominant hand towards a target stimulus appearing on
182 the left or right of the midline, while maintaining their head and eye position (Cameron,
183 Cheng, Chua, van Donkelaar, & Binsted, 2013) towards the central fixation point (which was
184 visible only prior to target onset). Participants were instructed to minimise both their reaction
185 times (RT) and movement times (MT). The participants were told that, in half of the trials, the
186 target would switch, either from left to right, or from right to left, and their task was to correct
187 their movement as rapidly as possible and point to the new target location if it switched. In
188 the multimodal conditions, participants were informed that in half of the trials the target would
189 change from visual to auditory or from auditory to visual, either on the same or a different
190 side. Participants were instructed to keep their index finger as still as possible at the end of
191 their pointing movement, until the end cue sounded.

192 Each trial began with a computer check of the position of the participant's index finger
193 and the orientation of their head, while the visual fixation stimulus was illuminated. If there
194 was any substantial deviation from the starting position for the finger (more than 2cm from
195 the start), or the straight-ahead orientation for the head (more than 7.5 degrees from straight
196 ahead), a warning tone was sounded. If the participant did not correct their finger or head
197 positions, the experimenter gave verbal prompts over an intercom. When both finger and
198 head were correctly positioned, the visual fixation was extinguished and there was a random
199 pre-trial interval with a uniform rectangular distribution between 1.5 and 3s. After the pre-trial

200 interval, the target stimulus was presented, and the participant made their movement. The
201 index finger position was analysed online, and in trials with a change in target location, the
202 change occurred as soon as the index finger's 3D velocity exceeded 10cm/s. This velocity
203 criterion was also used as the 'reaction time', both for trials with and without a target change.
204 Movement endpoint was the first sample with a 3D velocity below 5cm/s which was
205 maintained for at least 50ms. Participants were not able to touch the targets with their arm
206 outstretched. Changes in target location were achieved via a parallel port switch. The
207 stimulus waveform output was directed either to an LED or a speaker, depending on which
208 parallel port pin was active. Two seconds of position data were recorded, the end cue was
209 presented, and the visual fixation was re-illuminated before the start of the next trial.

210

211 *2.5.1 Calibration of body position, head orientation and EOG data*

212 Before the experiment, the position of the following body parts of each participant was
213 measured in the recording chamber: left and right index fingers, wrist, elbow, shoulder, and
214 neck, the top of the head (vertex) and between the two eyes. After the experiment, to
215 calibrate the head orientation with respect to the speaker locations, each participant was
216 asked to orient their head (using a laser pointer) to 7 locations, from -15 to +15 degrees right
217 of centre in 5 degree steps. A series of 7 tones was presented to cue the participants to
218 move, then fixate each location with the laser pointer. These data were used to calibrate the
219 raw head orientation data when computing gaze.

220 A similar procedure was used for EOG calibration. Guided by a series of LED
221 illuminations, participants made a series of 12 saccadic eye movements while keeping their
222 head oriented centrally. Each eye movement began at the central location, then fixated an
223 LED target at either -15, -10, -5, 5, 10, and 15 degrees right of centre for 1s, then returned to
224 the central location. These 12 saccades were used to calibrate the EOG data by regressing
225 the saccade-related change in mV EOG signal on the instructed change in eye position. The
226 slope of this regression was used to estimate eye position changes in degrees from the raw

227 EOG data. Supplementary Figure 4 shows mean horizontal and vertical EOG and gaze data
228 for the unimodal auditory experiment. All other experiments and conditions produced similar
229 EOG data.

230

231 **2.6 Data analysis**

232 *2.6.1 Preprocessing*

233 The experiments were run and the data analysis was performed using Matlab
234 (Mathworks, Natick, USA). All the programs are available from the first author's website
235 (<http://neurobiography.info>), and all raw data will be freely available there. Magnetic
236 interference from the loudspeakers and the metal hoop inside the chamber warped the
237 kinematic data. The warping of the measured space in the chamber was most severe at the
238 periphery, near the loudspeakers and furthest from the transmitter. Warping within the region
239 of most interest - the space traversed in the first few hundred milliseconds of movement
240 duration - was minimal. The kinematic data were unwarped during acquisition using a set of
241 multiple non-linear regression equations derived from a reference set of 500 known positions
242 measured inside the chamber (for a similar method, see Bryson, 1992).

243 Single time-point spikes due to random electromagnetic disturbances were removed
244 from the raw kinematic data in each of the three axes (x: near-far; y: left-right; z: down-up)
245 and replaced by linear interpolation from the adjacent points. The position data were then
246 upsampled to 240Hz then filtered with a 4th order zero-lag Butterworth low-pass filter with a
247 15Hz cut-off. Velocity and acceleration were calculated by simple differentiation, and a
248 number of kinematic parameters were extracted. In order to base our conclusions on the
249 maximum number of useful and valid trials, very broad ranges were used to accept valid
250 trials and to reject only very rare artefactual or clearly erroneous trials based on several
251 kinematic parameters. These parameters were decided upon based on experience and on
252 subjective visual exploration of the auditory only dataset, then programmed in for the
253 analysis of all three datasets. Similar parameters have been used elsewhere (e.g., for

254 minimum RT, Day & Brown, 2001; for maximum RT, the upper confidence limits of RTs in
255 Veerman et al., 2008 were about 650ms; See Supplementary Table 1 for further examples).
256 Exclusion criteria based on statistical thresholds can bias datasets (van Selst & Jolicoeur,
257 1994), so they were not used. The ranges were: RT (100-750ms); peak 3D acceleration
258 (100-16,000cm/s/s); time of peak 3D acceleration (0-500ms after RT); peak 3D velocity (25-
259 750cm/s/s); time of peak 3D velocity (16-1000ms after RT); peak 3D deceleration (-125 to -
260 16,000cm/s/s); time of peak 3D deceleration (33-1500ms after RT); MT (125-1500ms); path
261 length (10-100cm). Trials were also excluded if the maximum deviation of the head, eye, or
262 gaze was greater than 5 degrees in the first 200ms post-target presentation. This criterion
263 was relatively liberal for the EOG, but relatively conservative for the head and gaze
264 orientation data which were collected outside of the optimal motion tracking range for the
265 tallest participants.

266 Velocity data were aligned with respect to the RT (and, therefore, to the time that the
267 target changed, or would have changed, location). All timing parameters apart from RT were
268 measured relative to this point. For simplicity, and to improve sensitivity for the analyses of
269 most interest, data from left and right targets were combined by recoding the data as 'velocity
270 relative to the first target location', which involved inverting the y (lateral, left-right) axis data.
271 Similarly, the data from two left-handed participants were mirror-reversed across the midline.
272 We therefore refer to movements towards the ipsilateral and contralateral hemispaces: A
273 right-handed participant reaching towards the right target with their right hand is reaching into
274 ipsilateral hemisphere, while reaches with the right hand towards the left target are into
275 contralateral hemisphere. EOG data were downsampled to 240Hz, allowing alignment with
276 the kinematic data, then filtered with a 4th order Butterworth bandpass filter with cut-offs at 1
277 and 20Hz.

278 The development of analysis routines and parameters was based on the unimodal
279 auditory dataset. This raises the possibility that the analyses were biased towards the
280 properties or results of the auditory (exploratory) dataset. To counter this bias, the main

281 conclusions should therefore be based on the multimodal (confirmatory) dataset, in which the
282 unimodal auditory conditions were repeated. Differences between the unimodal auditory
283 experiment and the equivalent conditions in the visual or multimodal experiments should be
284 interpreted cautiously. To summarise: the auditory-only dataset was used to develop the
285 analytic routines, the visual-only dataset was used as a check to confirm whether and when
286 online corrections were present for the more standard unimodal visual targets, and the
287 multimodal dataset was used to explore the possibility of online control for targets that could
288 change modality as well as location.

289

290 *2.6.3 Measuring the time to initiate a movement correction*

291 The only dependent variable of interest was the latency to initiate a movement
292 correction. There are a number of methods for extracting this variable from movement data,
293 including visual inspection of trajectories (Alstermark *et al.*, 1984; Day & Brown, 2001; Hyde
294 & Wilson, 2013), use of an arbitrary distance threshold (Johnson, van Beers, & Haggard,
295 2002), examination of standard kinematic parameters or landmarks (e.g., Paulignan *et al.*,
296 1991a), classifying trials based on kinematic parameters (Pisella *et al.*, 2000), using
297 statistical criteria based on comparisons between corrected and uncorrected movements
298 using samples of trials (Baugh, Hoe, & Flanagan, 2012; Cressman, Cameron, Lam, Franks,
299 & Chua, 2010; Turrell, Bard, Fleury, Teasdale, & Martin, 1998) or samples of participants
300 (Aivar, Brenner, & Smeets, 2008; Glover, Miall, & Rushworth, 2005; Kerr, Fox, & Stein,
301 1994), or extrapolating from linear models fit to velocity curves (Veerman *et al.*, 2008).
302 Wijdenes *et al.* (2014) discussed and analysed the variety of methods of analysing correction
303 latencies available. Based on simulations of directional movement data with fixed and known
304 correction latency, they concluded that the model-fitting approaches applied to acceleration
305 data are best. Following reviewers' requests to justify our initial choice of method, we used
306 and developed a similar approach, comparing two methods of measuring the time to initiate a
307 movement correction, at three levels of analysis, and using three types of velocity data.

308 Unlike Wijdenes et al., (2014), we also examined a wide range of statistical thresholds for
309 determining correction latency, since choosing any threshold one was arbitrary. We tested
310 and optimised these eighteen methods on the unimodal auditory dataset, compared the
311 results with analysis of the visual dataset, then deployed them in the main, final analysis on
312 the visual and multimodal dataset.

313 The methods all used differences in the velocity between trials with a change in target
314 location versus trials with no change (i.e., correction velocity). In all cases, the modalities of
315 targets in the trials with and without changes were the same (i.e., trials with auditory targets
316 on the left followed by visual targets on the left were compared with trials with auditory
317 targets on the left followed by visual targets on the right. We examined both lateral correction
318 velocity (i.e., in the axis which defined the two target locations) as well as three-dimensional
319 correction velocities (derived from x-, y-, and z-axes), for two reasons – the pointing task was
320 three-dimensional, and the experimental axes were not perfectly aligned with the axes of
321 measurement. Table 1 and the following summarise the six main methods and three data
322 types.

323

324 [INSERT TABLE 1 ABOUT HERE]

325

326 • **Method 1: Group correction threshold.** Each participant's data were averaged to
327 produce a single correction velocity curve per condition. The statistics were then
328 based on the group mean of these trajectories.

329 • **Method 2: Participant correction threshold.** For each participant, the trials with a
330 change in target location are treated as a sample, and compared with the null
331 hypothesis of no correction. The one sample t-statistic reflects trajectory corrections
332 for each participant and condition.

- 333 • **Method 3: Trial correction threshold.** Each trial with a change in target location is
334 analysed separately. Each trial is compared with the sample of trials of a similar
335 condition in which the target did not change location. This produces a Z-statistic for
336 each sample, reflecting trajectory corrections for each trial.
- 337 • **Method 4: Group zero-crossing:** Same as Method 1, except using the zero-crossing
338 point extrapolated from the line joining the 25% and 75% points (relative to the
339 maximum correction velocity, Veerman et al., 2008) of the group mean correction
340 velocity for each condition.
- 341 • **Method 5: Participant zero-crossing.** Same as Method 4, except using the zero-
342 crossing points of the lines joining the 25% and 75% points of each participant's
343 mean correction velocity for each condition.
- 344 • **Method 6: Trial zero-crossing.** Same as Method 4, except using the zero-crossing
345 points of the lines joining the 25% and 75% points of each trial's correction velocity for
346 each condition.
- 347 • **Velocity a: Lateral correction velocity.** The velocity in the y-axis (left-right) on trials
348 with no change in target location subtracted from the same velocity on trials with a
349 change in target location.
- 350 • **Velocity b: 3D correction velocity.** The resultant velocity derived from velocities in
351 the x-, y-, and z-axes ($\text{resultant} = \sqrt{x^2 + y^2 + z^2}$) on trials with no change in target
352 location subtracted from trials with a change in target location.
- 353 • **Velocity c: Statistical components of correction velocity.** The correction velocity in
354 the x-, y-, and z-axes separately, expressed in confidence interval units (i.e., the
355 mean velocity divided by the SE or SD, then divided by the critical statistical value),
356 for trials with no change in target location subtracted from trials with a change in
357 target location. When the sum of statistical components is greater than 1, the
358 correction velocity is outside the (e.g., 95%) velocity confidence ellipsoid.

359

360 2.6.4 Statistical thresholds

361 Methods 1-3 all required that an arbitrary statistical threshold was chosen to define the
362 latency at which the correction has occurred (see Wijdenes et al., 2014 for detailed
363 discussion and simulations). Previous reports have used arbitrary thresholds of 2 standard
364 errors (e.g., Kerr *et al.*, 1994), confidence intervals (e.g., Turrell *et al.*, 1998), or 2 standard
365 deviations (Baugh *et al.*, 2012), or used the first significant time-point (e.g., Aiver *et al.*, 2008;
366 Cressman *et al.*, 2010; Glover *et al.*, 2005, see Supplementary Table 1). Methods 1 and 2
367 compared a sample of participants or trials with zero, so the statistical units were standard
368 errors (i.e., t-tests). Method 3 compared single trials with a sample mean, so the statistical
369 units were standard deviations (Z-tests). For each method, one-tailed tests were used as the
370 predictions were unidirectional: participants always moved from the initial target towards the
371 second target (cf Wijdenes et al., 2014). Choosing a statistical threshold was difficult. Initially,
372 we used a 1% probability cut-off (i.e., 2.68 standard errors for $n=13$, and 2.72 for $n=12$), as a
373 means of protecting against the increased false-positive rate in Method 1 (i.e., sequential
374 testing against zero for each possible correction latency over 50ms). However, there was no
375 reason to use this criterion for Methods 2 and 3, which extracted correction latencies from
376 individual participants or trials, then performed the final statistics at the group level.

377 Following an initial review, a systematic exploration of the statistical threshold for
378 determining correction latency was performed, since previous work has not examined the
379 effect of varying statistical threshold within a dataset (e.g., Wijdenes et al., 2014). Correction
380 velocities for Methods 1-3 were determined using 61 statistical thresholds, from 0.0 to 6.0
381 SE/SD in steps of 0.1. The minimum possible correction latency was set at 50ms, and in
382 each case, correction latency was defined as the first time-point in sequential testing that
383 exceeded the statistical threshold. For all 3 methods, as the statistical threshold increased
384 from 0 to 6, the mean correction latency increased from 50ms to over 300ms. Previous
385 studies have reported visuomotor correction latencies of 90ms or more, so we expected the

386 optimal methods to produce latencies in the 100-200ms range (Wijdenes, Brenner, &
387 Smeets, 2011; Archambault et al., 2015). During this exploratory analysis, one constraint we
388 thought important is that the method should be robust to small changes in the choice of
389 statistical threshold: if the sample size was increased, then the statistical threshold might
390 change (e.g., from 2.72 SE/SD to 2.68 for an increase from 12 to 13 participants).
391 Robustness to small changes in threshold seemed to be evident in how the correction
392 latencies changed as a function of statistical threshold: The more robust methods changed
393 smoothly with changes in threshold, the less robust, less smoothly. Finally, seeing these
394 threshold-latency curves, we reasoned that the peak of the curve may be informative: given a
395 method that is robust to small, arbitrary changes in statistical threshold, the statistical
396 threshold at which the greatest change in correction latency occurred should be informative
397 about the actual correction latency. Theoretically, the problem is one of using noisy data to
398 decide when a signal becomes non-zero. In the case of pointing velocity data, the signal is
399 likely Gaussian in shape. We reasoned that the point of maximum change in latency (x) as a
400 function of threshold (y) should correspond to the maximum slope of the velocity curve.
401 Numerical simulations confirmed this (see Supplementary Materials).

402 For subsequent analysis of the data, we chose the mean peak (across experiments
403 and conditions) of these latency-threshold curves (vertical lines in Figure 5a and 5b) as the
404 'optimal' threshold for each method (Table 1). This choice was arbitrary – given a sigmoid-
405 like increase in correction latency as a function of statistical threshold, we assumed that the
406 steepest part of the sigmoid may represent the best threshold to detect the correction
407 velocity signal above the noise. However, prompted by reviewers, we tested and verified this
408 method of choosing a statistical threshold using numerical simulation. See Supplementary
409 Materials and Supplementary Figures 1-3.

410 Using these three post-hoc criteria (correction latencies approximately between 100
411 and 200ms; correction latencies vary smoothly as a function of statistical threshold; peak
412 change in correction latency indicates optimal statistical threshold), Method 3 seemed the

413 most robust of those methods requiring an arbitrary statistical threshold (Table 1, Figures 3-
414 6). The optimum statistical threshold produced by this method was the least variable across
415 different experimental conditions (lowest coefficient of variation, CV, across conditions). The
416 threshold for Method 2 was the least robust and most variable across conditions. While these
417 criteria may have biased the subsequent analysis in general, they were applied equally to all
418 conditions, and all valid correction velocities ≥ 50 ms were analysed.

419 In the analyses reported below, Methods 1-3 used what we determined was the optimal
420 statistical threshold for each method (Table 1, right column, bold values; for Method 1, the
421 maximum statistical threshold across conditions was used since the latency-threshold curve
422 was not smooth for the lowest thresholds). For Methods 4-6, no statistical thresholds were
423 required, but some of the velocity data were unsuited to this analysis as they produced very
424 large variability in correction latencies (including negative latencies, cf Wijdenes et al., 2014).
425 Methods 4 and 5 (group and participant correction velocities, respectively) did not work well
426 for the sum of statistical components (4c and 5c), while Method 6 (trial correction velocities)
427 only worked for the lateral (y-axis) correction velocities (6a). Thus, results for fourteen of the
428 eighteen possible methods are reported. These failures of the model-fitting approaches likely
429 indicate that the correction signal was not strong enough in individual trials, or even
430 participants, to produce a meaningful correction velocity with these three-dimensional
431 velocity measures. Rather than being a weakness of these methods, it may be a strength -
432 the correction signal is primarily, or entirely in the lateral velocity component, and adding
433 other directional velocity components merely adds noise to this signal.

434

435 **3 Results**

436 The only dependent variable of relevance to the aims of the study was the latency to
437 correct movements following a change in target location. Additional analyses, for
438 completeness, to ensure the experimental conditions were comparable, and to answer
439 reviewers' comments, are reported as Supplementary Materials. Participants were able to

440 correct their reaching-to-point movements in all conditions: visual, auditory, visual-auditory,
441 and auditory-visual targets. Figure 2 shows the group mean (across participant means)
442 position (upper panel), velocity (middle panel), and acceleration (lower panel) curves for the
443 four unimodal auditory conditions. Red and blue curves show trials where the target
444 remained stationary, and magenta and cyan where it switched sides at movement onset. The
445 data were pooled across ipsilateral and contralateral targets to create displacements and
446 velocities relative to the first target location (i.e., towards the ipsilateral side if the first target
447 was in ipsilateral space). The velocity data were then analysed according to the fourteen
448 different methods described above (Figure 3).

449

450 **3.1 Exclusion of data**

451 A mean(\pm SD) of 5.54(6.63), 7.62(11.9), and 16.1(18.5) trials were removed from the
452 visual-only, auditory-only, and multimodal experiments (7.83, 6.22, & 9.64%, respectively).
453 Of a total of 4955 trials, 364 (7.92%) were removed. 93 (2.02%) were removed for RTs below
454 100ms, 9 (0.19%) on the peak acceleration criteria, 8 (0.17%) on peak velocity, 5 (0.11%) on
455 peak deceleration, 72 (1.57%) on movement time, 22 (0.48%) on path length, 26 (0.57%) on
456 eye position, 1 (0.02%) on head position, 118 (2.57%) on eye velocity (i.e., saccades or EOG
457 artefacts), and 10 (0.22%) on gaze velocity (i.e., combined eye/head shifts not otherwise
458 detected). Supplementary Figure 4 shows the mean eye position and gaze orientation across
459 participants. While the position of the eye in the head, as measured by EOG, was quite
460 stable, most participants seemed to make quite large head rotations while reaching-to-point,
461 affecting the overall gaze orientation. The head orientation data are, however, less reliable,
462 particularly in tall participants, due to the distance of the head receiver from the Polhemus
463 transmitter. Importantly, however, the mean head and gaze orientation was not significantly
464 different from zero throughout the whole trial, and by the time the mean head orientation had
465 changed more than a few degrees, the reaching trajectory corrections had already begun
466 (i.e., within 300ms).

467

468 **3.2 Methods of determining correction latency**

469 The first main aim of the study was to evaluate different methods of determining
470 correction latency for three-dimensional velocity data, particularly where those methods
471 require an arbitrary statistical threshold for determining when the movement correction
472 begins. Six methods and three types of velocity data were examined.

473

474 *3.2.1 Method 1: Group correction thresholds*

475 Since only one statistical comparison was performed on the group-level data, the
476 statistical t-value plotted against the correction latency gave a single curve for each
477 experiment and condition (e.g., Figure 5A, solid black line). Correction latencies for all six
478 experiments and conditions (Unimodal auditory A, and visual V; multimodal VV, AA, VA, and
479 AV) were between 167 and 242ms, with means for Methods 1a, 1b, and 1c of 202ms,
480 206ms, and 195ms respectively. Data for the unimodal auditory condition are shown in
481 Figure 6.

482

483 *3.2.2 Methods 2 and 3: Participant and trial correction thresholds*

484 Methods 2 and 3 depended critically upon an arbitrary statistical threshold chosen to
485 define the correction latency. Method 2 used the mean and variability on a participant-by-
486 participant basis, while Method 3 used a trial-by-trial analysis. The resulting participant
487 means were then analysed at the group level. In order to choose a threshold, we
488 systematically varied the threshold and examined the statistical main effects and interactions
489 across all experiments and conditions. To illustrate the problem of choosing a threshold,
490 Figure 4 shows the effect of the initial threshold (x-axis, from 0 to 6, SE for Method 2, SD for
491 Method 3) on the resulting statistical effects in the multimodal experiment (y-axis, t-values).
492 This figure shows a 'dance of the t-values' (cf Cumming, 2012) – how the primary statistical
493 effects of interest (y-axes) change as a function of the initial (and arbitrary) statistical

494 threshold (x-axis). Figures 4a-4c show that for most thresholds, there were no main effects of
495 the initial target modality (most lines are between the critical t-value criteria), but that for
496 some methods and some threshold ranges, significant main effects or interactions can be
497 found. Indeed, for at least one threshold level per method, four of six methods produced a
498 significant main effect of the initial target modality, four of the final target modality, and all six
499 produced significant interactions between initial and final target modalities.

500 Given the arbitrary choice of statistical threshold, we required a more constrained
501 method of extracting the correction latency for Methods 2 and 3. Exploring the data, we
502 plotted the mean correction latency (Figure 5a) and the change in mean correction latency
503 (Figure 5b) as a function of the initial statistical threshold. For Methods 2a, 2b, and 2c the
504 resulting curve was quite erratic with no clear peak (Figure 5b shows Method 2a, broken
505 lines). By contrast, for Methods 3a, 3b, and 3c, the curve was smoother, and contained a
506 single clear peak (Figure 5b shows Method 3a, solid light grey line).

507

508 *3.2.3 Methods 4-6: Group, participant, and trial zero-crossings*

509 One clear advantage of a model-fitting approach is that there is no need for an arbitrary
510 statistical threshold to determine the correction latency. Instead, in the three model-fitting
511 methods used here, a straight line is fit to two points on the correction velocity curve, the first
512 at 25% of the maximum correction velocity, the second at the 75% point (Figure 3d, Veerman
513 *et al.*, 2008). This analysis was performed on the three correction velocity types at group,
514 participant, and trial levels. Only five of the nine possible methods succeeded in measuring
515 correction latency, with the lateral correction velocity (a) providing seemingly the most robust
516 inputs for Methods 4, 5, and 6.

517

518 *3.2.4 Comparison of methods of determining correction latency*

519 Our first aim was to compare different methods of determining correction latency based
520 on real velocity data for three-dimensional pointing movements (cf Wijdenes *et al.*, 2014).

521 Figure 6 shows the performance of fourteen different methods for determining correction
522 latency in the unimodal auditory dataset, which was used to optimise the analysis routines.
523 From these data, while Method 2 (participant-level data) produced the lowest estimates of
524 correction latency (Figure 6a, across-condition range=110-230ms), they were also the most
525 variable (SD and CV, Figures 6b, 6c; cf Veerman et al., 2008). To measure the correlation
526 between different methods, r-values determined across participants within each condition
527 were Z-transformed, then averaged across experimental conditions. The mean correlations
528 between correction latency measurements determined by Method 2 and by the other
529 methods (Figures 6d and 6e) were the lowest. The opposite pattern applied to Methods 4-6
530 (zero-crossing, model-fitting): the highest correction latencies (range=182-271ms), but the
531 least variable and most highly-correlated with other methods. Method 3 produced mean
532 (range=130-218ms) and variable correction latencies between those of Methods 2 and
533 Methods 4-6. In summary, Method 2 produced short but variable correction latencies which
534 may provide an estimate of the earliest movement corrections; Methods 4-6 provided longer
535 but less variable latencies; Method 3 may provide a compromise between these two options.

536

537 **3.3 Effects of target modality**

538 Our second aim was to determine correction latency for our auditory, in comparison to
539 our visual target objects. We decided that Method 6a – fitting a straight-line model to lateral
540 correction velocity data on a trial-by-trial basis was the best method. Similar conclusions
541 were reached by others (Wijdenes et al., 2014). We applied this method to analyse the
542 effects of target modality in the three experiments. The unimodal auditory experiment
543 resulted in similar correction latencies (mean \pm SD=244 \pm 49.2ms) to the unimodal visual
544 experiment (241 \pm 50.1ms, $t(12)=-0.274$, $p=.789$). For the same comparison in the multimodal
545 experiment, however, the unimodal auditory corrections (233 \pm 62.0ms) were initiated
546 significantly earlier than the unimodal visual (255 \pm 53.1ms, $t(11)=2.29$, $p=.043$).

547 A factorial analysis with the variables initial target modality (auditory, visual), and final
548 target modality (auditory, visual) revealed a significant effect of initial target modality,
549 $t(11)=6.68$, $p<.001$, with corrections away from auditory targets initiated 35.6 ± 18.5 ms faster
550 than away from visual targets. There was no significant effect of final target modality ($p=.08$)
551 or interaction between initial and final target modality ($p=.166$).

552

553 **3.4 Relationship between target localisability, target salience, and correction latency**

554 Our auditory and visual targets were matched for location and timing, were driven by
555 the same signal, and were both clearly suprathreshold, but they differed in physical size. We
556 did not explicitly equate auditory and visual targets for localisability, detectability, or salience
557 prior to running the experiments. Rather, in pilot testing we ensured that the auditory stimulus
558 was maximally localisable, and used a relatively dim LED for the visual stimulus. Equating
559 different stimulus attributes is very difficult within vision alone (Veerman et al., 2008, p220),
560 and is perhaps even more so between modalities. Can differences in correction latency
561 between auditory and visual stimuli be explained by differences in localisability of the
562 targets? Analysis of the means across participants' mean endpoint constant errors, both in
563 distance and direction, suggested that, indeed, visual targets were localised between 2.0-
564 2.6mm and 0.1-0.2 degrees better (closer to the target) across experimental conditions and
565 endpoint measurements than auditory. These small differences (approximately half of the
566 LED's diameter; 0.5% of the total movement length; 1% of the angle between targets) were
567 significant in 3 of 5 variables examined (uncorrected p-values .012 to .033). However,
568 arguing against the possibility that localisability *explains* correction latency, only one of 20
569 correlations performed between endpoint error and correction latency (on different measures,
570 both trial-by-trial, and using participants' means) was significant (uncorrected $p=.035$,
571 correlation between angular error relative to the head, and correction latencies measured
572 with Method 3a). Note, too that minimising endpoint error was not emphasised to the
573 participants, that no instructions were given about exactly how to point (e.g., "position the

574 fingertip on a line between the eye/head and the target”), that the greatest magnetic
575 distortions in the positional data were at the movement endpoint, and that many stimulus and
576 task manipulations affect endpoint accuracy in purely visual experiments (see
577 Supplementary Materials).

578 Since it is not clear how auditory and visual conspicuity (or salience) are experimentally
579 to be equated (cf Veerman et al., 2008), we used the proxies for conspicuity that Veerman et
580 al. (2008) used – correction velocity slope – and that Cameron et al., (2013) used – peak
581 correction velocity magnitude. Correction magnitude systematically affects the calculation of
582 correction latency (Wijdenes et al., 2014). Comparing the unimodal (visual only and auditory
583 only) experiments, there was no significant difference in either correction velocity slope
584 ($t(12)=1.01$, $p=.334$) or magnitude ($t(12)=1.59$, $p=.138$). For correction velocity slopes in the
585 multimodal experiment, there was no significant main effect of initial ($F(1,11)=1.42$, $p=.258$),
586 or of final target modality ($F(1,11)=0.002$, $p=.968$), and no significant interaction between
587 these factors ($F(1,11)=0.20$, $p=.663$). For peak correction velocity magnitudes, there was
588 also no significant main effect of initial ($F(1,11)=0.87$, $p=.371$) or final target modality
589 ($F(1,11)=1.21$, $p=.294$), while the interaction between these variables only showed a trend
590 ($F(1,11)=4.48$, $p=.058$) in which corrections within a modality resulted in non-significantly
591 greater peak correction velocities (visual mean \pm SD=103 \pm 28.9cm/s; auditory=103 \pm 24.5cm/s)
592 than corrections between modalities (visual-then-auditory=99.6 \pm 29.2cm/s, auditory-then-
593 visual=96.0 \pm 26.3cm.s), this trend is in the same direction as the non-significant interaction in
594 correction latencies reported above. These results argue against there being differences
595 between the localisability, detectability, salience, or conspicuity of the auditory and visual,
596 and to a lesser extent the multimodal, target objects in our experiment (Veerman et al., 2008;
597 Cameron et al., 2013).

598

599 **4 Discussion**

600 We examined the latency for making online corrections in reaching-to-point movements
601 towards visual, auditory, visual-then-auditory, and auditory-then-visual targets. Across
602 experiments, conditions, and methods, significant movement corrections were evident at
603 110-271ms after the change in target location. These correction latencies were mostly much
604 lower than the mean reaction times for initiating the movements (238-296ms).

605 Our first finding is that fitting a straight-line model (Methods 4-6) to velocity data
606 provides better (lower variability) correction latencies than using a statistical threshold or
607 sequential statistical testing (Methods 1-3). Second, with the model-fitting approaches,
608 examining the component of velocity only in the direction of the target jump is more likely to
609 be successful than using all three directional components (e.g., Methods 4a, 5a, and 6a
610 here). Third, if using a statistical threshold, then the choice of threshold to determine
611 correction latency (whether in SE, SD, or confidence interval units, at the single trial or single
612 participant level) can have dramatic and unpredictable effects on the outcome of subsequent
613 statistical testing (Figure 4). Without examining a wide range of possible statistical thresholds
614 systematically, any choice of threshold is arbitrary and potentially misleading. Similar
615 conclusions were reached by Wijdenes et al., (2014), who decided, based on simulations of
616 lateral movement corrections with known onset, magnitude, and intensity, that fitting straight-
617 line models to acceleration data was optimal. Our work extends their findings by looking
618 systematically at 61 different statistical thresholds, three different types of velocity in three
619 dimensions, and testing real data from three-dimensional movements.

620 Our second novel finding is that the online control of reaching-to-point movements can
621 be just as effective, in latency and magnitude, for auditory as for visual targets (e.g., Boyer *et*
622 *al.*, 2013; Veerman *et al.*, 2008), at least in a unimodal context, where target modality was
623 fixed. Our third finding is that the latency of online corrections of movements to targets that
624 can change modality as well as location (i.e., movement corrections in a multimodal context)
625 depends upon the initial target modality – corrections away from auditory targets are initiated
626 earlier than corrections away from visual targets. These conclusions must be tempered by

627 the fact that we chose only one specific kind of auditory and visual target. Different stimulus
628 attributes may result in different correction latencies (Veerman et al., 2008). Indeed, it may
629 be that whether the participant (or just the relevant parts of their nervous system) treats the
630 two targets as the same is critical for initiating rapid corrections. This is discussed further
631 below.

632

633 **4.1 Latency of corrections to targets presented unimodally**

634 Across the two unimodal experiments, correction latencies did not differ significantly,
635 however within the context of the multimodal experiment, latencies to correct movements to
636 purely auditory targets were significantly shorter (22ms, using the model-fitting Method 6a)
637 than to purely visual targets. This difference is slightly lower than might be expected based
638 on the difference in initial processing time (i.e., RT) for visual over auditory targets: Mean
639 (\pm SD) RTs to visual targets were 39.1 ± 37.8 ms longer than for auditory targets (see
640 Supplementary Materials). Very similar differences in RT and the latency of auditory and
641 visual signals in superior parietal lobe were reported by Molholm, Sehatpour, Mehta,
642 Shpaner, Gomez-Ramirez, Ortigue, Dyke, Schwartz, & Foxe (2006). Why this RT advantage
643 did not translate into a similar advantage in correction latency in the purely unimodal
644 experiments is unclear. One caveat is that the auditory dataset was used for exploration and
645 to optimise the analysis methods, so comparisons between this and other datasets need to
646 be made cautiously (i.e., are potentially biased by 'double-dipping').

647 In general, the correction latencies reported here are slightly longer than those reported
648 in other similar research (e.g., a minimum of 90ms in Paulignan et al., 1991a; \sim 130-200ms
649 across participants in Veerman et al., 2008; see Supplementary Table 9 for further details).
650 Two relatively trivial factors, and one likely more important factor may contribute to this: First,
651 our kinematic data were smoothed with a 15Hz low-pass cut-off filter, which is higher than
652 some other similar studies (5Hz, Boyer et al., 2013; 8Hz, Paulignan et al., 1991a) - lower-
653 frequency cut-offs smooth the data more. In exploratory work, repeating the analyses with a

654 5Hz cut-off resulted in correction latencies an average of 26ms shorter. Second, our
655 kinematic data were recorded, and target location changed, only at 60Hz, meaning that our
656 correction latencies are over-estimated by an average of 8.33ms (half a sample). Finally, and
657 likely most important, our targets were presented in three-dimensional space rather than, for
658 example, on a graphics tablet, computer screen, or other flat surface (e.g., Veerman et al.,
659 2008). This complicates the movements performed, requiring different muscle groupings. It
660 also means that, since the axes of measurement and axes of movement were not perfectly
661 coregistered, components of the movement corrections may have occurred in the x-, y-, and
662 z-axes rather than purely in the lateral y-axis. We attempted to overcome this limitation by
663 examining both lateral correction velocity (Methods 1a-6a, Figure 3a, similar to much
664 previous research), as well as resultant 3D velocity (Methods 1b-6b, Figure 3b), and the
665 individual statistical components of correction velocities (Methods 1c-6c, Figure 3c). Overall,
666 different sampling and analysis parameters may have produced perhaps 20-30ms greater
667 correction latencies, but by far the largest contribution to estimation of correction latency is
668 the method and statistical threshold chosen (Figures 5-6). For example, Leonard, Gritsenko,
669 Ouckama, & Stapley (2011) studied reaching and pointing movements in 3D space,
670 measuring correction latencies in 2D, using Method 3 (individual trials) with a 1SD threshold.
671 They reported correction latencies of around 180-190ms. Using the same statistical threshold
672 and method, our latencies for the unimodal auditory experiment are 128ms for the 1D data
673 (Method 3a), and 121ms for the 3D data (Method 3b).

674

675 **4.2 Latency of corrections to targets presented multimodally**

676 The primary aim of this research was to assess whether the online control of reaching-
677 to-point movements might be multimodal or supramodal, or whether it is, at least in
678 significant part, unimodal. This aim was operationalised by hypothesising that, if online
679 control is substantially unimodal, then having to switch between visual and auditory targets,

680 or auditory and visual targets, should incur some cost relative to changing movement
681 trajectories between targets of the same modality. This hypothesis was not supported.

682 The first target modality had a greater effect on correction latencies (36ms, $d=1.9$) than
683 the second target modality (14ms, $d=0.56$). For the first target, it was easier for participants
684 to correct their movements away from the auditory targets than away from the visual targets.
685 This finding might be explained by hypothesising that reaching movements to visual targets
686 are more 'locked on' to their target; that movements towards initially visual targets are more
687 ballistic and less amenable to online control than movements to auditory targets; or that
688 reaching and pointing movements are generated and controlled using predominantly visual
689 representations of target location, that it is easier to select, maintain, and acquire visual
690 targets than auditory targets. This could be due to a relative imprecision in auditory
691 localization relative to visual localization. This possibility can be tested in future research by
692 systematically manipulating the relative localisability of targets across modalities, and
693 therefore the precision of movements towards them (e.g., Izawa & Shadmehr, 2008).
694 Examining the latency to correct movements towards auditory targets in congenitally blind,
695 recently blind, and blindfolded participants may also shed light on this question. Finally,
696 under an attentional account, auditory localization may be more dependent on focusing
697 attention than visual localization which may be more automatic. Therefore, any attentional
698 switch cost may be larger when one has to switch from visual to auditory than vice versa.

699

700 **4.3 Target detectability, localisability, and number**

701 Following Veerman et al. (2008), and Cameron et al. (2013), we measured the
702 correction velocity slope and peak correction velocity as proxies for the conspicuity (i.e.,
703 detectability, or salience) of our different modality targets, and found no significant effects of
704 target modality, apart from a trend towards corrections within a modality having higher peak
705 correction velocity than between modalities. Without another available measure of
706 conspicuity between modalities, these data suggest that our targets were well-matched, and

707 that the significant differences in correction latencies that we reported are not due to
708 differences in conspicuity. Nevertheless, we can of course not claim that all kinds of auditory
709 and visual targets will produce the same patterns of results. Within vision alone, the
710 particular attributes of the stimuli determine to a great extent the correction latency (Veerman
711 et al., 2008), and this is almost certainly true for auditory, proprioceptive, and tactile targets
712 as well (Cameron & López-Moliner, 2015).

713 Previous visual studies of online control have used, for example, different initial and
714 final target size and colour (Day & Brown, 2000), or compared correction latencies between
715 objects differing on numerous stimulus attributes and task relevance (Aivar et al., 2008,
716 2015), or examined corrections away from visual targets that were presented shortly after
717 imperative auditory movement cues (Wijdenes et al., 2011). Similarly, many previous studies
718 have used two or more discrete target objects, LEDs, or illuminated locations on a flat
719 screen, with target illumination switching instantaneously between the two stimuli, rather than
720 moving a single, physical object (Supplementary Table 9; see Day & Lyon, 2001, for a
721 counter-example; and a brief review in Sarlegna & Muthi, 2015). Thus, the present study, as
722 well as most previous studies, of online control implicitly assumes that participants perceive
723 continuity (i.e., apparent motion) between the (illuminated) target locations; that the same,
724 single object is apparently moved. This is the assumption of unity (for examples in
725 multimodal perception, see Vatakis & Spence, 2007). If the online control of movements
726 depends on the assumption of unity, this concern about whether online corrections are the
727 same for a single, moved object versus two sequentially presented objects or locations
728 applies equally to unimodal and multimodal studies alike, and represents an important
729 question for future research to address.

730

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738

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830 **7 Figure Legends**

831 **Figure 1. Experimental apparatus.** Participants sat at the centre of a 90cm radius metal
832 hoop (grey arc) supporting three loudspeakers (filled trapeziums) and three LEDs (filled
833 circles) at the centre, and 7.5 degrees to the left and right of the participant's midline.
834 Participants rested their hand on a starting board (grey rectangle), keeping their index finger
835 in a 'start' location (filled circle). Participants wore a tracker on their index finger and vertex
836 (solid squares), and a pair of goggles supporting a laser pointer (filled rectangles).

837

838 **Figure 2. Mean (\pm SE) position (a), velocity (b), and acceleration curves (c).** Data show
839 reaching-to-point movements towards an auditory target 7.5 degrees on the ipsilateral (red,
840 "Ipsi-ipsi") or contralateral side of the midline (blue, "Contra-contra"). The magenta curves
841 show movements initially directed to an ipsilateral target, and corrected after movement
842 onset to the contralateral target (magenta, "Ipsi-contra"), and cyan curves show movements
843 initially directed to the contralateral, then corrected to the ipsilateral target (cyan, "Contra-
844 ipsi").

845

846 **Figure 3. Methods of measuring latency to correct a reaching movement.** Each panel
847 shows a correction velocity curve with correction latency derived in different ways. (a) The
848 mean (thick black line) lateral (ipsi-to-contra) velocity data are used to find the first point
849 where the lower confidence limit (lower thin black line) of the correction curve is above zero
850 (broken red line). The vertical red line indicates the correction latency. The confidence
851 interval around the mean correction velocity is set according to an arbitrary choice of
852 statistical threshold (e.g., 2.18 standard errors from the mean would be a 95% CI for a
853 single-sample *t*-test with 13 participants). This procedure is used in Methods 1a, 1b, 2a
854 (example data shown), and 2b. (b) A similar procedure can be used on single trials with a
855 change in target location, by comparing a single correction velocity curve to the mean \pm SD
856 velocity on trials without a change in target location. Correction latency is determined when

857 *the correction velocity exceeds the confidence limit (broken red line). This procedure is used*
858 *in Methods 3a and 3b (example data shown). (c) A 'confidence ellipsoid' is calculated using*
859 *the x (blue), y (green), and z (magenta) statistical components of the correction velocities.*
860 *The statistical components of the x, y, and z correction velocities (i.e., the mean of each*
861 *component of velocity, divided by the SD or SE of the components, then divided by the*
862 *critical statistical threshold value (e.g., $t(12)=2.18$ for a 95% confidence ellipsoid). Correction*
863 *latency is determined by the first point where the sum of the three components is greater*
864 *than one (i.e., outside the confidence ellipsoid, broken red line). This procedure is used in*
865 *Methods 1c, 2c, and 3c. (d) A model-fitting approach is used to fit a straight line to the first*
866 *points along the correction velocity curve from panel a which are greater than or equal to*
867 *25% (25% V_{max} , lower solid horizontal red line) and 75% (75% V_{max} , upper solid horizontal*
868 *red line) of the maximal correction velocity. The straight line model is then extrapolated back*
869 *to find the zero crossing with the x-axis (broken red line). The zero-crossing is the correction*
870 *latency (vertical red line).*

871

872 **Figure 4. Effect of statistical thresholds for determining correction latency on main**
873 **effects of and interactions between initial and final target modality in the multimodal**
874 **experiment.** *Each panel shows the statistical threshold used to determine the correction*
875 *latency on the x-axis (from 0 to 6, corresponding to the number of standard errors from the*
876 *mean for participant mean data (black lines, Methods 2a, 2b, 2c), and standard deviations for*
877 *trial-by-trial data (grey lines, Methods 3a, 3b, 3c). The y-axes show (a) the statistical main*
878 *effects (t-values) of the initial target modality, (b) the final target modality, (c) and the*
879 *interaction between initial and final target modality (c). Horizontal red lines show the critical t-*
880 *values for significant effects (5%, two-tailed). Depending on both the method and the*
881 *statistical threshold chosen, both main effects and their interaction can be found 'significant'*
882 *or 'not significant'.*

883

884 **Figure 5. Effect of statistical thresholds on correction latency in the unimodal auditory**
885 **experiment.** Each plot shows the statistical threshold on the x-axis and (a) the resulting
886 mean lateral correction latency or (b) the change in mean lateral correction latency on the y-
887 axis. Method 1a (black line) produced only a single group correction latency (the first valid
888 (i.e. ≥ 50 ms) correction latency occurred with the threshold indicated by the vertical line,
889 hence the curve is truncated; the curve is smoothed for display purposes), but Methods 2a
890 (broken grey line) and 3a (sold grey line) produced a different correction latency for each
891 participant and condition (Method 2a) or participant, condition, and trial (Method 3a). The
892 data in panel b show that the group mean correction latency changes more smoothly as a
893 function of the statistical threshold for Method 3a than for Method 2a. Method 3a produces a
894 smoother curve (panel a), and a single large peak (panel b), while Method 2a produces a
895 less smooth curve and multiple peaks. The point of maximal change in correction latency as
896 a function of statistical threshold (i.e., the steepest part of Figure 5A – horizontal lines on
897 each curve) was taken as the optimal statistical threshold (vertical lines) to use for
898 determining correction latency. See Supplementary Materials for a numerical simulation and
899 validation of this approach.

900

901 **Figure 6. Descriptive and diagnostic statistics for fourteen methods of determining**
902 **correction latency.** Each plot shows the group mean ($\pm 95\%$ CI where available) of
903 correction latency statistics from each of fourteen methods, extracted from the correction
904 velocity data. (a) Mean correction latency for all 14 methods. (b) Standard deviation (SD) of
905 correction latency across participants for Methods 2, 3, 5, and 6. (c) Coefficient of variation
906 (CV) correction latency (SD/mean). (d) Correlations between different procedures of
907 estimating correction latency within each Method (e.g., the correlations for Method 2a are
908 with Method 2b, and 2c). (e) Correlations between different procedures and methods (e.g.,
909 the correlations for Method 2a are with Methods 3a, 3b, 3c, 5a, 5b, and 6a). Correlations are

910 *expressed as Z-values after Fisher's r-to-Z transformation, to allow valid use of parametric*
911 *statistics.*

912 **Table 1: Performance of nine methods of determining correction latency for visual,**
913 **auditory, visual-then-auditory, and auditory-then-visual targets using statistical**
914 **thresholds**

Method	Level	Velocity	Units	Optimal threshold per condition						Mean (SD) (CV)
				VV	AA	VV	VA	AV	AA	
1a	Group	lateral	SE	1.70	2.56	2.42	2.82	2.06	1.87	2.24 (0.432) (0.193)
1b	Group	3D-r	SE	2.58	4.47	2.68	1.13	1.91	1.28	2.34 (1.22) (0.522)
1c	Group	3D-c	SE	2.73	1.88	1.85	2.92	3.07	1.68	2.36 (0.618) (0.262)
2a	Subject	lateral	SE	1.2	0.9	2.2	1.1	2.3	1.6	1.55 (0.589) (0.380)
2b	Subject	3D-r	SE	1.1	2.0	2.5	1.4	2.1	2.2	1.88 (0.527) (0.280)
2c	Subject	3D-c	SE	0.6	0.5	1.0	0.9	1.5	0.8	0.883 (0.354) (0.401)
3a	Trial	lateral	SD	1.2	1.0	1.4	1.2	1.3	1.2	1.22 (0.133) (0.109)
3b	Trial	3D-r	SD	1.2	1.5	1.9	1.6	1.3	1.0	1.42 (0.319) (0.225)
3c	Trial	3D-c	SD	1.9	2.1	2.5	1.7	2.0	2.4	2.10 (0.303) (0.144)

915 Data shown are the optimal thresholds for determining correction latencies in SE (standard
916 error) or SD (standard deviation) statistical units. CV: Coefficient of variation=SD/mean. VV:
917 Visual-visual; AA: Auditory-auditory; VA: Visual-auditory; AV: Auditory-visual; Mod. 1: Main
918 effect of initial target modality; Mod. 2: Main effect of final target modality. 3D-r: Resultant 3D
919 velocity= $\sqrt{x^2+y^2+z^2}$; 3D-c: Statistical components of velocity in x-, y-, and z-axes.

920 **Table 2. Fourteen measures of correction latency for visual, auditory, visual-then-**
 921 **auditory, and auditory-then-visual targets**

Method	Unimodal			Multimodal						
	VV	AA	diff (p)	VV	AA	VA	AV	Mod 1 diff (p)	Mod 2 diff (p)	Interaction diff (p)
1a	179	192	-13.0	217	217	221	183	19.0	-19.0	15.0
1b	200	204	-4.00	217	179	242	196	42.0	-4.00	-21.0
1c	183	192	-9.00	217	204	217	171	31.5	-18.5	-14.5
2a	143 (46.5)	121 (37.6)	21.6 (.10)	110 (37.3)	118 (57.9)	181 (180)	126 (50.1)	23.8 (.47)	-31.5 (.31)	-40.0 (.17)
2b	167 (42.6)	153 (71.5)	14.6 (.47)	121 (57.1)	138 (65.2)	176 (55.5)	167 (45.6)	-4.08 (.84)	-12.8 (.44)	-41.6 (.028)
2c	230 (59.3)	214 (53.4)	16.5 (.51)	112 (59.3)	155 (80.4)	155 (53.0)	139 (39.3)	-7.73 (.71)	-34.6 (.032)	-7.73 (.62)
3a	152 (21.4)	175 (54.4)	-22.4 (.10)	137 (28.4)	130 (34.2)	164 (53.9)	145 (31.1)	13.3 (.16)	-5.74 (.44)	-21.0 (.13)
3b	186 (32.9)	211 (34.6)	25.7 (.06)	168 (35.8)	183 (43.1)	218 (59.8)	202 (54.4)	0.76 (.95)	-15.5 (.24)	-34.3 (.10)
3c	175 (35.0)	191 (43.0)	-16.5 (.28)	149 (26.8)	145 (38.8)	208 (65.8)	177 (50.6)	17.8 (.24)	-13.8 (.25)	-45.3 (.013)
4a	206	210	-3.47	223	199	222	186	30.1	-6.31	6.54
4b	219	223	-3.71	227	189	235	213	30.3	7.91	-16.5
5a	229 (46.6)	223 (44.8)	5.60 (.56)	248 (46.8)	215 (52.3)	244 (47.2)	211 (46.8)	32.8 (.005)	0.02 (.99)	3.98 (.50)
5b	247 (49.7)	246 (46.9)	1.03 (.92)	254 (52.8)	227 (53.7)	260 (48.2)	237 (48.0)	25.0 (.002)	1.73 (.78)	-7.63 (.15)
6a	246 (50.1)	248 (49.2)	-2.67 (.79)	255 (53.1)	233 (62.0)	275 (55.2)	225 (57.2)	35.6 (<.001)	-14.0 (.08)	-5.99 (.17)

922 Data shown are the mean (SD) correction latencies in ms using six methods (Table 1). p and
 923 ANOVA columns (Mod 1, Mod 2, Interaction) show p-values for t-test and ANOVA terms
 924 respectively. diff=effect size in ms (visual-auditory or unimodal-multimodal). Bold font
 925 indicates significant effects (p<.05, uncorrected). VV: Visual-visual; AA: Auditory-auditory;
 926 VA: Visual-auditory; AV: Auditory-visual; Mod. 1: Main effect of initial target modality; Mod. 2:
 927 Main effect of final target modality.