1	Sex biases in kin shoaling and dispersal in a cichlid fish
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19	
20	Abstract
21	Animal dispersal is associated with diverse costs and benefits that vary among
22	individuals based on phenotype and ecological conditions. For example, females may
23	disperse when males benefit more from defending territories in familiar environments.

24	Similarly, size differences in dispersal propensity may occur when dispersal costs are
25	size-dependent. When individuals do disperse, they may adopt behavioral strategies that
26	minimize dispersal costs. Dispersing fish, for example, may travel within shoals to reduce
27	predation risks. Further, kin shoaling may augment inclusive fitness by reducing
28	predation of relatives. However, studies are lacking on the role of kin shoaling in
29	dispersal. We explored how sex and size influence dispersal and kin shoaling in the
30	cichlid Neolamprologus caudopunctatus. We microsatellite genotyped over 900
31	individuals from two populations separated by a potential dispersal barrier and
32	documented patterns of population structure, migration and within-shoal relatedness.
33	Genetic differentiation across the barrier was greater for smaller than larger fish,
34	suggesting larger fish had dispersed longer distances. Females exhibited weaker genetic
35	differentiation and 11 times higher migration rates than males, indicating longer-distance
36	female-biased dispersal. Small females frequently shoaled with siblings, possibly
37	offsetting dispersal costs associated with higher predation risks. In contrast, small males
38	appeared to avoid kin shoaling, possibly to avoid local resource competition. In
39	summary, long-distance dispersal in N. caudopunctatus appears to be female-biased and
40	kin-based shoaling by small females may represent a behavioral adaptation that reduces
41	dispersal costs. Our study appears to be the first to provide evidence that sex differences
42	in dispersal influence sex differences in kin shoaling.
43	
44	Keywords Animal movements $\cdot$ Fish shoals $\cdot$ Grouping behavior $\cdot$ Kin selection $\cdot$

45 Population genetics

## 47 Introduction

48 Dispersal by individuals from their natal grounds to their breeding areas is a common 49 feature of animal life (Greenwood 1980) and is associated with significant benefits and 50 costs to dispersers (Handley and Perrin 2007; Bonte et al. 2012). By dispersing, animals 51 may avoid the negative effects of inbreeding (e.g. Nagy et al. 2013; Banks and 52 Lindenmayer 2014) and increase access to resources such as territories or mates (Bowler 53 and Benton 2005). In addition, they may increase inclusive fitness, for example, by 54 relaxing local resource competition among relatives (reviewed in Handley and Perrin 55 2007). However, diverse costs of dispersal also exist, including increased risks, and loss 56 of time and energy (Bonte et al. 2012). For example, since dispersers often travel 57 solitarily across unfamiliar and unsuitable habitat, a major cost is an increased risk of 58 predation (e.g. Hiddink et al. 2002; Yoder et al. 2004). 59 The decision to disperse may accordingly depend on a tradeoff between these 60 costs and benefits, which in turn is predicted to vary with the ecological conditions and 61 individual phenotypes (Bowler and Benton 2005). For example, size-differences in 62 dispersal strategies may arise when only larger individual have sufficient energy reserves 63 to be able to move large distances or be competitive in new environments (e.g. 64 Gundersen et al. 2002). Similarly, local mate- and resource competition may lead to sex-65 differences in dispersal strategies. In birds, for example, females typically disperse, as 66 males gain more from being able to defend a familiar territory in their natal habitat (Greenwood 1980). In contrast, male-biased dispersal occurs more often in polygynous 67 68 mammals as a result of females preferring immigrant males to avoid inbreeding (Handley

69	and Perrin 2007). Sex-biases in dispersal strategies in turn have important consequences
70	for the genetic structure of populations (e.g. Sheridan et al. 2010; Bisol et al. 2012).
71	Unlike the well-studied birds and mammals, the causes and consequences of
72	dispersal strategies in fish are less well-understood and often less clear-cut. For example,
73	in the Atlantic salmon (Salmo salar), temporal and spatial variation in sex ratios appear
74	to favor dispersal of different sexes (Consuegra and García de Leániz 2007).
75	Additionally, the biology of fish differs markedly from that of birds and mammals, which
76	may further affect dispersal strategies. For example, body size may affect the swimming
77	performance of fish during dispersal (e.g. Stobutzki and Bellwood 1997; Fisher et al.
78	2000) and their competitive ability in new environments (Rowland 1989; Koops and
79	Grant 1993). In addition, larger fish may suffer less predation from gape-limited fish
80	predators (Scharf et al. 2000). Therefore, as fish are known to continually grow
81	throughout their lives (Helfman et al. 1997), dispersal may be delayed until a threshold
82	body size is reached.
83	When individuals do disperse, they are expected to adopt behavioral adaptations
84	that minimize dispersal costs. For example, kin-based grouping is common in fish (e.g.
85	Arnold 2000; Frommen and Bakker 2004; Gerlach and Lysiak 2006) and may be
86	beneficial during dispersal by reducing stress associated with interacting with unfamiliar
87	conspecifics (e.g. Gerlach et al. 2007) or augmenting inclusive fitness via a reduction in
88	predation of relatives (Hatchwell 2010). When only one sex disperses, sex differences in
89	grouping behavior may be expected, with kin-based grouping favored in the dispersing
90	sex. For species that are non-cooperative or do not live in family groups, these kin-based
91	groups would be expected to dissolve prior to breeding to reduce kin conflict for

92 resources (Hatchwell 2010). Studies are lacking however, on the relationship between93 dispersal and kin-grouping.

94 We characterized patterns of dispersal and kin-shoaling in *Neolamprologus* 95 *caudopunctatus*, a predominately colonial cichlid from Lake Tanganyika. Dispersal 96 strategies within the genus *Neolamprologus* appear to be highly variable, with some 97 species displaying male-biased dispersal (e.g. N. pulcher, Stiver et al. 2004), while in 98 other species females disperse (N. multifasciatus, Schradin and Lamprecht 2000). N. 99 *caudopunctatus* is a socially monogamous and sexually monomorphic substrate breeder 100 (Schaedelin et al. 2013). Individuals which are not currently breeding form large shoals 101 that are often located above breeding colonies. In contrast, breeding pairs occupy the 102 substrate, search for a nesting site together and construct nesting cavities by excavating 103 sand under stones (Ochi and Yanagisawa 1999). Nesting cavities and free swimming fry 104 are defended by both parents over a period of several weeks until fry are independent 105 (Ochi and Yanagisawa 1999). Despite both sexes being involved in nest site selection and 106 defense, males are larger than females (Schaedelin et al, unpublished data), suggesting 107 that they are more effective in territorial defense, as occurs in other fish (Rowland 1989; 108 Koops and Grant 1993). In support of this, males display a higher attack rate towards 109 territorial intruders than do females (Ochi and Yanagisawa 1999).

We genetically sampled over 900 breeders and shoaling fish in two distinct populations. These populations are separated by a potential dispersal barrier consisting of extended areas of sandy substrate where rocks are scarce, a known dispersal barrier for this species (Koblmueller et al. 2006). We first quantified genetic population structure within and between populations to detect evidence of sex-biased dispersal. The higher 115 aggressiveness of males in this species suggests that they may benefit more from 116 defending a territory in a familiar environment, thus favoring female dispersal 117 (Greenwood 1980). Our second aim was to test whether kin-based shoaling occurred in 118 this species as a possible strategy to reduce the costs of long-distance dispersal. Finally, 119 we determined whether any tendency to shoal with kin was size- or sex-dependent. As 120 smaller fish presumably have reduced swimming performance (Fisher et al. 2000), are 121 less competitive (Rowland 1989; Koops and Grant 1993; Aubin-Horth et al. 2007) and 122 are younger (Helfman et al. 1997), we predicted that genetic signatures of dispersal are 123 more likely to be detected amongst larger individuals.

124

### 125 Materials and methods

126 *Field work* 

127 The study was carried out in October and November 2008 in Kasakalawe Point (S 08° 128 46' 46.6" E 31° 04' 44.4") on the southern shore of Lake Tanganyika, southwest of 129 Mpulungu, Zambia (hereafter referred to as the 'mainland site'). Clusters of N. 130 *caudopunctatus* defending potential nesting cavities were identified along the shore at 131 depths of 5 to 12m. We identified and numbered 165 breeding cavities defended by pairs 132 in an area of approximately 220 x 250 meters by SCUBA diving. To map the whole 133 population underwater (including the location of breeders and shoal fish, see below), we 134 constructed a coordinate system out of sisal rope and measured nest and shoal distances 135 to this coordinate plane using measuring tape. In addition to the main study site, we also sampled individuals inhabiting the waters around Mbita Island (S 08° 45' 20.2" E 31° 05' 136 137 27.0"), approximately 3 km from the main site (hereafter referred to as the 'island site').

We collected samples along 150 meters of the shore at a depth of 1.5 to 3m. We did not
map the population at the island site. However, sampling at this site allowed us to explore
sex-biased dispersal over a relatively larger spatial scale. Both study sites comprise
typical *N. caudopunctatus* habitat, with a sandy substrate interspersed with many rocks
(Konings 1998).

143 During the study period we detected no breeding activity, although many pairs at 144 the mainland site actively defended a breeding cavity. We refer to these paired, territorial 145 individuals as breeders, whereas individuals in shoals are referred to as shoal fish. 146 Breeders regularly join shoals that are located close to their breeding cavities for short 147 periods of time, possibly for feeding (Schaedelin, personal observation). We genetically 148 sampled breeders by capturing pairs defending a breeding cavity with monofilament nets 149 and measured their standard length. No anaesthesia was used. We then clipped 150 approximately 7.5mm of the dorsal fin in situ. Fin clips were stored in individual water-151 filled tubes during the diving session and transferred to Eppendorf tubes with absolute 152 ethanol once back on land. We assured that we did not resample individuals by checking 153 for evidence of our fin-clipping. Individuals were sexed by ventral inspection of their 154 genital papillae.

Shoal fish were captured and sampled in a similar fashion. We defined a shoal as an aggregation of fish where fish were separated by a maximum of three body lengths from each other (e.g. Pitcher and Parrish 1993). We estimated the approximate size of each shoal at the time of sampling. The proportion of fish captured per shoal varied widely depending on the size of the shoal. For small shoals (e.g. less than 20 individuals) we captured the majority of the fish, while for large shoals (e.g. greater than 100

161	individuals) we typically captured less than fifty percent of the fish. As we found no pairs
162	defending a breeding cavity at the island site, we only captured shoal fish from this
163	population. However, the lack of breeders captured at the island site is unlikely to bias
164	our analyses on size-dependent dispersal in this species as we still captured many fish of
165	potential breeding size at this site (see Results). In total, we sampled 117 breeders and
166	580 shoal fish originating from 35 shoals along the mainland site (mean number of fish
167	sampled from each shoal = $16.6 \pm 12.5$ fish; range = $1 - 63$ ) and 224 shoal fish from 9
168	shoals from the island (mean number of fish sampled from each shoal = $24.9 \pm 9.3$ fish;
169	range = 6 - 33).

### 171 *Genetic analyses*

172 DNA extraction of tissue samples was conducted using a QIAGEN DNeasy Blood and 173 Tissue Kit. All adults were genotyped at 11 microsatellite loci previously developed for 174 other species including NP773 and UNH002 (Schliewen et al. 2001), Pzeb3 (van Oppen 175 et al. 1997), TmoM5 and TmoM13 (Zardoya et al. 1996), UME003 (Parker and Kornfield 176 1996), UNH106, UNH130 and UNH154 (Lee and Kocher 1996), and UNH908 and 177 UNH1009 (Carleton et al. 2002). PCR and fragment analyses were conducted as 178 described in Schaedelin et al. (2013). Descriptive statistics were calculated for all 11 179 microsatellite loci (Online Resource 1), including the number of alleles (N<sub>A</sub>), observed 180 heterozygosity (H<sub>0</sub>), allelic diversity (H<sub>s</sub>) and the inbreeding coefficient (F<sub>IS</sub>) using 181 FSTAT 2.9.3 (Goudet 1995). We tested for Hardy-Weinberg equilibrium of all loci in 182 Cervus 3.0 (Kalinowski et al. 2007).

183	A principal aim of our study was to document differences in genetic structure
184	based on cichlid body size, which can provide information on dispersal strategies. Due to
185	difficulties in conducting many of the statistical analyses described below using size as a
186	continuous variable, we assigned individuals to size categories. We used the median
187	standard length of each sex as a threshold for each size category to ensure similar sample
188	sizes between size categories. We therefore classified females as small when their
189	standard length was less than 4.0cm, while small males were those shorter than 4.4cm.
190	These values were slightly larger than the minimum observed size of breeders at the
191	mainland population (males = $3.8$ cm, females = $3.7$ cm). Fish above these thresholds were
192	therefore likely able to defend territories and breed.
193	
194	Population structure
195	We quantified the sex- and size-specific genetic structure of the study populations using
196	several complementary approaches (i.e. analyses of molecular variance, Bayesian cluster
197	analyses, population tree reconstruction, coalescence modeling and the characterization
198	of shoal-sharing by siblings). Using multiple analytical methods allowed us to
199	comprehensively and robustly track sex- and size-differences in genetic structure at the
200	levels of the shoal and population. Mantel tests were carried out to test for isolation by
201	distance within the mainland population (for which we had geographic distance data),
202	using GENETIX (Belkhir et al. 2004), by correlating a measure of genetic distance,
203	$\frac{F_{ST}}{1-F_{ST}}$ (Rousset 1997), with geographic distance between individual shoals of fish.
204	We conducted analyses of molecular variance (AMOVA) in ARLEQUIN 3.5.1.2
205	(Excoffier and Lischer 2010) to determine whether sex and size differences in dispersal

affect the partitioning of genetic variation. In total, four AMOVAs were conducted. We divided our total sample into sex and size specific data sets (small males, large males, small females, large females) and defined the two study sites as groups and shoals within the study sites as populations. For twelve shoals, we only had genetic samples for six or fewer individuals. In contrast, the remainder of shoals (n = 32) contained a minimum of eleven individuals. We therefore reduced sample size biases by excluding the twelve small shoals from our analyses.

213 Although an AMOVA can detect population genetic differentiation, it provides no 214 information on migration rates, which are expected to be higher for the dispersing sex. 215 Therefore, we analyzed male and female data sets using an isolation with migration 216 model (IM: Hey and Nielsen 2007), which assumes a split in a parent population at some 217 time in the past, but with continued gene flow (migration) between the two resulting 218 daughter populations. Since the model is based on the coalescent, bi-directional migration 219 rates and effective population sizes can be simulated simultaneously. Thus, we used the 220 software IMa (Hey and Nielsen 2007) to test for differences in migration rates between 221 the male and female populations. IMa uses a MCMC Bayesian approach to simulate the 222 posterior distributions of migration rates between island and mainland populations ( $m_1$ 223 and  $m_2$ ) and effective population sizes of the parental ( $\theta_A$ ) and daughter populations ( $\theta_1$ 224 and  $\theta_2$ , respectively). The population migration rate for each sex (i.e. the effective rate, 225 per generation, at which genes come into a population) can then be calculated as migration rate<sub>1</sub> =  $\frac{\theta_1 m_1}{2}$ . A series of preliminary runs were first executed to determine the 226 227 prior parameter distributions to be used in subsequent runs. We then conducted three 228 independent runs for each sex, using 20 Markov chains with geometric heating for five

229	million generations after an initial burn-in period of one million generations. All runs
230	converged and estimates of the model parameters were the same for each run. Due to the
231	computationally demanding nature of these analyses, we did not conduct similar analyses
232	separating sexes into size classes.
233	We used STRUCTURE 2.3.3 (Pritchard et al. 2000) to estimate the number of
234	genetic clusters, $K$ , in our data set. As we were interested in sex differences, we analyzed
235	each sex separately. Each individual was assigned to a cluster with probability, $q$ . Five
236	independent analyses were performed for $1 \le K \le 10$ using a Markov chain of 600,000
237	iterations, discarding the first 100,000 as burn-in. To avoid potential biases, we assumed
238	the admixture model without including prior knowledge of the origin of samples.
239	Correlated allele frequencies were also assumed. The most likely number of clusters in
240	the data set was estimated using two methods. We first calculated $\Delta K$ following Evanno
241	et al. (2005), which estimates the most likely number of clusters based on the rate of
242	change of log-likelihood probabilities for each K. As this method cannot detect the best K
243	when K=1, we also estimated the number of clusters following Pritchard et al. (2000),
244	which uses Bayes' rule for the log-likelihood probabilities.
245	Although STRUCTURE can infer the number of genetic clusters within a dataset,
246	it cannot provide information on the relationships among populations. We therefore
247	created population neighbor joining phylogenetic trees using the $\frac{F_{ST}}{1-F_{ST}}$ distance
248	calculated in ARLEQUIN to ascertain whether shoals were grouped by geographic
249	origin. The matrix was then imported into PHYLIP (Felsenstein 2005) to create
250	population phylogenies using the neighbor-joining method from Saitou and Nei (1987).
251	

252 Genetic similarity

We calculated genetic similarity between all shoal fish and breeders within each population, following Mathieu et al. (1990). Here, similarity was estimated as the probability that a given pair will produce homozygous offspring (Phm). For each locus (1), Phm is equal to:

257 
$$Phm_{xy}(l) = \frac{(s_{ac} + s_{ad} + s_{bc} + s_{bd})}{4}$$
, where s<sub>ij</sub> equals 1 if alleles i and j are the same, and 0

258 otherwise. Across all loci, a weighted average was used:  $Phm_{xy} = \frac{\sum_{l} \frac{1}{p_{l}} Phm_{xy}(l)}{\sum_{l} \frac{1}{p_{l}}}$ , where

p<sub>1</sub> is the probability of an individual being homozygous by chance at locus l. Therefore,
Phm<sub>xy</sub> will be closer to 1 for more genetically similar individuals, who are thus more
likely to produce homozygous offspring. Belkhir et al. (2002) showed that this index of
genetic similarity is a superior alternative index when the number of loci used is
relatively low (i.e. compared to genetic similarity estimates calculated following Queller
and Goodnight 1989 and; Lynch and Ritland 1999)

265 We tested whether size and sex differences affected mean population-wide 266 relatedness by conducting analyses separately for shoal fish and breeders. Following 267 Taylor et al. (2003) we first created separate relatedness matrices for each sex. To avoid 268 pseudoreplication we averaged the relatedness of each individual with all other 269 individuals of the same sex. We then performed a two-sample permutation test using 270 10,000 randomizations using RUNDOM (Jadwiszczack 2002) on these averaged data to 271 test whether the difference in male and female relatedness differed significantly from 272 differences obtained by randomly assigning individuals to each sex.

273	We were also interested in whether sex and size differences affected the
274	relatedness of individuals within versus among shoals (i.e. whether fish join shoals of
275	related individuals). We therefore separated the shoals into the two size classes (i.e. small
276	and large fish) to explore whether small fish were more related within a shoal than larger
277	fish. We only conducted these additional analyses for the mainland population owing to
278	small sample sizes for the island shoals once the data were separated into size and sex
279	classes. We compared the mean relatedness of all fish within a shoal with relatedness
280	among individuals from different shoals using PERM (Duchesne et al. 2006). PERM uses
281	matrices of a pairwise relatedness statistic ("Sxy" - which corresponds to Phm in this
282	study) and calculates the sum of all Sxy values (i.e. Phm sums) within each group (shoals
283	in our study). These values are then compared with a distribution of Sxy sums generated
284	from randomly assigning individuals to shoals. One thousand randomizations were used.
285	Although the above analyses provide information on the average relatedness of
286	fish within shoals, they provide no information on the actual number of siblings that
287	shoal together. We therefore identified probable siblings based on Phm values to explore
288	what proportion of shoals consisted of siblings. To identify the minimum Phm of siblings
289	(and the maximum Phm of non-siblings) we first generated 500 offspring from matings
290	between 50 randomly selected male and female territorial pairs using HybridLab (Nielsen
291	et al. 2006). This generated a relatedness matrix of known siblings and non-siblings and
292	their corresponding Phm values. Based on these data (Online Resource 2) we
293	conservatively assigned two individuals as siblings when their Phm was greater or equal
294	to 0.57 and they were within the same size class. This value corresponds to the first
295	quartile for all siblings (i.e., 75% of all siblings have Phm values above 0.57) and is

296	greater than the maximum Phm value generated for non-siblings (maximum Phm for non-
297	siblings = 0.55). We then calculated sex and size differences in 1) the proportion of
298	individuals within a shoal that are siblings (pooling all shoals of a given sex and size
299	class and using a chi-square to test for group differences), 2) the proportion of siblings
300	that were captured within the same shoal as opposed to being captured in different shoals
301	and 3) the mean geographic distance between all sibling pairs (calculated based on the
302	position of each shoal on our coordinate system at the time of capture).
303	All non-genetic statistical analyses were conducted using SPSS 17.0 (SPSS,
304	Chicago, Illinois, USA). Non-parametric tests were used when the assumption of data
305	homoscedasticity was not met. All data are presented as mean $\pm$ SD.
306	
307	Results
308	Shoal characteristics
309	Shoal size was highly variable, ranging from five individuals to over 200 individuals. At
310	the mainland population, where both breeder and shoal fish were sampled, breeders were
311	larger than shoal fish (mean standard length: breeders = $4.6 \pm 0.4$ cm, shoal fish = $4.2 \pm$
312	0.6 cm; Mann-Whitney U = 15,349, n = 623, P < 0.001). The percentage of fish within
313	shoals classified as "large" was highly variable, ranging from 0-100% (mean = $57\pm36\%$ )
314	at the mainland population and 53-100% (mean = $88\pm17\%$ ) at the island population.
315	Among shoals, mean male and female size of each shoal was positively correlated (r =
316	0.851, n = 22, $P < 0.001$ ). Finally, high variation in sex ratios between shoals was
317	observed, ranging from 20-100% females per shoal (mean = $55\pm18\%$ , n = 23) at the
318	mainland population and 21-67% at the island population (mean = $52\pm15\%$ , n = 8).

320 Sex-size differences in overall population structure

321	We found no evidence of isolation by distance within the mainland population for either
322	males (mantel test: $Z = 70.43$ , $P = 0.336$ ) or females (mantel test: $Z = 453.37$ , $P = 0.590$ ).
323	Similarly, we found no evidence of isolation by distance when separating individuals by
324	sex and size (mantel test: small males - $Z = 53.61$ , $P = 0.852$ ; large males - $Z = 151.44$ , $P$
325	= 0.188; small females - $Z = 152.49$ , $P = 0.282$ ; large females - $Z = 73.64$ , $P = 0.608$ ).
326	Finally, we found no difference in the average relatedness between individuals within
327	sexes for either shoal fish or breeders, even when separate analyses were conducted for
328	small and large fish (Table 1).
329	AMOVA revealed that in both sexes the majority of genetic variation
330	(approximately 98%) occurred among individuals within shoals (Table 2). However,
331	genetic variance among shoals within study sites was over twice as high for small
332	females (0.83%) than for small males (0.36%: chi-squared test - $\chi^2 = 18.56$ , d.f. = 1, $P <$
333	0.001), but similar for large females and males (0.63% and 0.55%, respectively: chi-
334	squared test - $\chi^2 = 0.54$ , d.f. = 1, $P = 0.461$ ; Table 2). Furthermore, the only among-shoal
335	variance value (i.e. $F_{SC}$ ) that was greater than zero was for small females. Less than 2.5%
336	of variation was explained by differences between the mainland and island populations,
337	although the male populations were slightly more differentiated than the female
338	populations. For both sexes, genetic differentiation between the two study sites was
339	higher for small than large individuals (i.e. based on $F_{ST}$ values in Table 2). An analysis
340	of migration rates using a coalescent IM model, revealed that migration from the island to
341	the shore was four times higher for females than for males and 11 times higher for

females than for males migrating from the shore to the island (based on  $\frac{\theta_i m_i}{2}$  values in

343	Table 3).
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344	Bayesian population assignment using STRUCTURE supported the one-cluster
345	(K = 1) hypothesis as most likely for females, whereas $K = 2$ returned the highest log-
346	likelihood for males (Fig. 1, Online Resource 3). Among males, the two clusters
347	approximately defined the mainland and island populations (mean probability of
348	assignment $q$ (cluster 1/mainland) = 0.219 ± 0.244, $q$ (cluster 2/island) = 0.852 ± 0.201;
349	$F_{1,387} = 654.36$ , $P < 0.001$ ). Forty seven percent (183/389) of the males had a $q$ value of
350	less than 0.90, and 3.9% (8/206) of individuals with $q$ >0.90 could not be assigned to the
351	population from which they were sampled. To further investigate population genetic
352	structuring of the shoals, we created sex-specific neighbor-joining phylogenies based on
353	our microsatellite data (Fig. 2). The trees supported the findings above, showing a clear
354	division between the mainland and island populations for males, but not for females.
355	
356	Sex- and size-dependent relatedness among and within shoals
357	We found sex and size differences in relatedness within shoals. Within-shoal relatedness
358	was significantly higher than between-shoal relatedness only for small females (Table 4).
359	In contrast, neither large females, nor large or small males were more closely related to

360 fish of their size and sex class within their shoal than that expected by random shoaling.

In further support of kin-shoaling in small females, we found that small females
were more likely than individuals of other classes to be captured together with at least
one sibling (proportion of sibling pairs captured within same shoal: small females – 18/58
(31.0%), large females – 3/40 (7.5%), small males – 7/72 (9.7%), large males – 5/37

365	(13.5%); logistic regression – sex – Wald = 8.56, d.f. = 1, $P = 0.003$ , size – Wald = 0.36,
366	d.f. = 1, $P = 0.551$ , sex*size – Wald = 5.24, d.f. = 1, $P = 0.022$ ). Furthermore, small
367	females shared shoals with more siblings than did fish in the other classes (proportion of
368	siblings within shoals: small females - 16% (24/150), small males - 7% (11/165), large
369	females – 2% (3/146), large males - 4% (7/160); chi-squared test – $\chi^2 = 25.00$ , d.f. = 3, <i>P</i>
370	< 0.001). Short distance dispersal also involved sex- and size-specific differences (mean
371	sibling-sibling distance: small females – 72.8 $\pm$ 62.9m, large females – 90.7 $\pm$ 62.9m,
372	small males $-110.5 \pm 71.5$ m, large males $-74.2 \pm 71.0$ m; ANOVA: sex $-F_{1,203} = 1.11$ ,
373	$P = 0.293$ , size $-F_{1,203} = 0.84$ , $P = 0.362$ , sex*size $-F_{1,203} = 7.28$ , $P = 0.008$ ). Small male
374	siblings were geographically more dispersed among shoals than small females ( $F_{1,128} =$
375	9.93, $P = 0.002$ ) and large males ( $F_{1,107} = 6.32$ , $P = 0.013$ ). Adult male siblings were not
376	more dispersed than adult female siblings ( $F_{1,75} = 0.99$ , $P = 0.323$ ).

## 378 Discussion

379 We have undertaken a genetic study of over 900 Neolamprologus caudopunctatus across 380 a localized scale of up to 3 km in Lake Tanganyika to determine factors affecting 381 population structure. Our results highlight several sex- and size-specific characteristics 382 that provide a detailed overview of the dispersal strategies adopted by this species. 383 Fundamentally, we found that females dispersed longer distances than males. There was 384 strong agreement among the several complementary analyses that we implemented to 385 detect this sex difference. Weaker genetic differentiation among females between the two 386 study sites (Table 2, Figure 1, 2), coupled with 11 times higher female migration rates 387 between the two populations (Table 3), support the conclusion that longer-distance

388 dispersal is female-biased. Interpopulational genetic differentiation was greater for small 389 than large females, suggesting that larger females had dispersed longer distances than 390 smaller females. However, the small amount of genetic variation that could be explained 391 by population differences suggests that significant gene flow still occurs between the two 392 populations despite the presence of the potential dispersal barrier. Although the barrier 393 presumably consisted predominately of a sandy substrate with a scarcity of rocks, the 394 possibility remains that suitable habitat exists within this area that acts as a stepping stone 395 for fish dispersing between the sites. At the more localized scale of the mainland 396 population, small female siblings tended to shoal together (Table 4). This pattern 397 occurred despite females not being more related to each other at the population level than 398 were males (Table 1). Longer-distance dispersal was less frequent among males, as 399 shown by the pronounced genetic differentiation across our study area, and by lower male 400 migration rates. Yet, at shorter distances, small male siblings were sampled in more 401 geographically separated shoals than other size and sex classes, implying a greater 402 propensity for kin avoidance among small males. The contrast in dispersal strategies of 403 males and females, despite the existence of shoals of mixed sex and size, implies that 404 shoaling behavior in *N. caudopunctatus* may be characterized by constant fission and 405 fusion.

406

407 Size and sex differences in dispersal

408 Our data suggest that it is more beneficial for females to disperse than to remain on their 409 natal grounds. This may, in part, be related to the potential advantage of males remaining 410 in a familiar area to facilitate territory acquisition and defense, as also occurs in birds 411 (Greenwood 1980). By dispersing, females may also reduce the probability of breeding 412 with related individuals and have access to new territories (Greenwood 1980; Handley 413 and Perrin 2007). However, dispersal between the mainland and island site is also 414 presumably costly. N. caudopunctatus are exposed to many fish predators, including 415 species that prey on fry (e.g. Telmatochromis vittatus and N. tetracanthus; Ochi and 416 Yanagisawa 1999) and adults (e.g. Lepidiolamprologus elongatus; Schaedelin, personal 417 observation). These costs may be augmented by the scarcity of rocks along the dispersal 418 barrier between the two sites, making it difficult for individuals to seek shelter from 419 predators. Dispersing individuals are therefore expected to adopt strategies that minimize 420 these costs. For example, as our data suggest, individuals may benefit from dispersing 421 over long distances only when they have reached a relatively larger size, to minimise 422 predation from gape-limited fish predators and to benefit from improved swimming 423 performance (Fisher et al. 2000; Scharf et al. 2000).

424 In addition, kin-based shoaling may provide females with indirect benefits 425 associated with decreased predation of relatives during dispersal. Recent research has 426 revealed that shoaling is often kin-based (reviewed in Hatchwell 2010), the advantages 427 of which are diverse (Ward and Hart 2003). Shoaling with kin may reduce stress, thus 428 facilitating more rapid growth (Gerlach et al. 2007). Further, Piyapong et al. (2011) found 429 that kin-based shoals of juvenile Trinidadian guppies (Poecilia reticulate) were more 430 likely to form in habitats with high predation pressure. This suggests that individuals may 431 gain indirect benefits by decreasing the predation of close relatives. Several studies have 432 also detected sex differences in kin-shoaling (e.g. Griffiths and Magurran 1998; Arnold 433 2000; Ruhl and McRobert 2005; Gerlach and Lysiak 2006). In the Trinidadian guppy, for

example, only females preferred to shoal with familiar kin (Griffiths and Magurran
1998), which may reflect a trade-off between the advantages of shoaling with familiar
individuals and moving among shoals in search of mates. However, there has been a lack
of previous studies to suggest that sex and size differences in kin shoaling may arise from
differences in dispersal strategies.

439 In contrast to females, dispersal of male N. caudopunctatus appears to be more 440 localised. This reduced dispersal propensity may result in higher competition between 441 sibling males for territories or mates (e.g. Frommen et al. 2007). Selection may therefore 442 favour stronger kin avoidance for males at this scale (West et al. 2002; Handley and 443 Perrin 2007). Although large male siblings were not more dispersed than large female 444 siblings, we found that small male siblings were more dispersed than the three other size 445 and sex classes, suggesting that they actively avoid shoaling with kin. Given the longer-446 distance dispersal of large individuals, the greater localised dispersal of small males is 447 curious and difficult to explain with our current data, thus warranting further research. 448 However the results imply that the benefits of male kin avoidance may be size-dependent. 449 For example, if body size is correlated with age in this species (e.g. Helfman et al. 1997) 450 and young males search for mates in shoals, kin avoidance by small males may reduce 451 local mate competition among siblings. In addition, males may display lower fidelity to 452 individual shoals to permit mate searching in multiple shoals (e.g. Griffiths and Magurran 453 1998).

454 Our study highlights the complexity of dispersal strategies (short distance male-455 biased dispersal, but longer-distance female dispersal) and stresses the importance of 456 exploring dispersal at multiple spatial scales. To our knowledge our study system is the 457 first to demonstrate sex-biased dispersal and kin-based shoaling in tandem. Our findings458 may therefore provide new insights into the benefits of kin-grouping in animals.

459

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473	

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621	Table 1 Mean	relatedness	of individuals	within age	and sex o	categories. F	values r	efer to

Class	Class 1	Class 2	Class 1 Phm ± SD	Class 1 Phm ± SD	Р
Shoal	Females	Males	0.266±0.03 (264)	0.266±0.04 (231)	0.930
Shoal	Small females	Small males	0.269±0.03 (128)	0.270±0.03 (126)	0.702
Shoal	Large females	Large males	0.267±0.03 (136)	0.267±0.04 (105)	0.894
Breeder	Females	Males	0.265±0.03 (64)	0.267±0.04 (60)	0.709

622 the statistical significance of the difference between two classes in within-class Phm.

**Table 2** AMOVA table for *N. caudopunctatus*, separated by sex and size, outlining variation among shoals and sites.

Small males								
Source of variation	d.f.	Sum of squares	Variance components	Percentage of variation	Fixat	ion indices	P value	
Among sites	1	15.8/3	0.08573	2.40	F <sub>ST</sub>	0.02755	<0.001	
Among shoals within sites	22	80.446	0.01272	0.36	$F_{SC}$	0.00365	0.522	
Within shoals	322	1119.190	3.47575	97.25	F <sub>CT</sub>	0.02399	< 0.001	
Total	345	1215.509	3.57420					
Large males								
Among sites	1	9.782	0.04769	1.38	F <sub>ST</sub>	0.01926	0.003	
Among shoals within sites	13	48.237	0.01897	0.55	F <sub>SC</sub>	0.00556	0.242	
Within shoals	243	824.876	3.39455	98.07	F <sub>CT</sub>	0.01378	0.003	
Total	257	882.895	3.46122					
			Small females					
Among sites	1	13.466	0.07292	2.02	F <sub>ST</sub>	0.02847	< 0.001	
Among shoals within sites	17	69.112	0.02982	0.83	F <sub>SC</sub>	0.00843	0.01	
Within shoals	347	1216.545	3.50589	97.15	F <sub>CT</sub>	0.02021	< 0.001	
Total	365	1299.123	3.60863					
Large females								
Among sites	1	8.837	0.04190	1.18	F <sub>ST</sub>	0.01817	0.004	
Among shoals within sites	17	64.273	0.02240	0.63	F <sub>SC</sub>	0.00641	0.111	
Within shoals	247	858.188	3.47444	98.18	F <sub>CT</sub>	0.01184	< 0.001	
Total	265	931.297	3.53875		01			

**Table 3** Posterior parameter estimates for male and female cichlids after simulation using coalescent IM modelling. Refer to Methods

for explanation of model parameters. The population migration rate for each sex and population was calculated following	$\frac{\theta_1 m_1}{2}$ .
	for explanation of model parameters. The population migration rate for each sex and population was calculated following

	Posterior model parameters					Demograp	hic conversion
	θ1	θ2	θ <sub>A</sub>	<i>m</i> <sub>1</sub>	<i>m</i> <sub>2</sub>	$\frac{\theta_1 m_1}{2}$	$\frac{\theta_2 m_2}{2}$
Males	0.089±0.041	0.259±0.089	0.988±0.011	100.97±42.27	23.48±20.31	4.49	3.04
Females	0.240±0.046	0.594±0.134	0.990±0.010	147.70±31.46	112.55±30.16	17.72	33.43

- 627 **Table 4** Difference between observed intrashoal Phm sum and that expected due to
- for a random mixing of individuals. Data are presented as mean  $\pm$  SD. When Phm sums are

629 larger than expected, individuals within shoals are more related than expected by chance.

Class	Observed intrashoal Phm sum	Expected intrashoal Phm sum	N	Р
Small females	180.1	172.6±2.7	11	0.001
Large females	114.4	113.6±2.2	12	0.322
Small males	116.0	117.0±2.2	17	0.625
Large males	104.6	107.6±2.4	16	0.859
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## 645 Figure legends

646 **Fig. 1** Posterior population assignment probabilities among a) female and b) male *N*.

647 *caudopunctatus*. Cluster analyses were conducted in STRUCTURE (K=1 for females,

648 K=2 for males). Each vertical line represents one individual and the proportion of white

649 to grey reflects the probability of belonging to each of the two clusters.

650

**Fig. 2** Neighbour-joining trees for a) female and b) male *N. caudopunctatus* sampled in

the two populations. Data are based on 11 microsatellite loci. Only shoals for which we

653 genotyped a minimum of eight individuals of one sex were included in the analysis. Each

654 circle represents a single shoal. Empty circles represent shoals captured at the island site

and filled circles represent shoals from the mainland site.

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Fig. 1





Fig. 2