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# The impact of fungicide treatment and Integrated Pest Management on barley yields

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# 1 The impact of fungicide treatment and Integrated Pest Management on barley yields: 2 analysis of a long term field trials database 3 Stacia Stetkiewicz<sup>1, 2, 3,4,</sup>, Fiona J. Burnett<sup>1</sup>, Richard A. Ennos<sup>3</sup>, Cairistiona F.E. Topp<sup>1\*</sup> 4 5 6 7 8 <sup>1</sup> Crops and Soil Systems, Scotland's Rural College, Peter Wilson Building, King's Buildings, W. Mains Road, Edinburgh EH9 3JG <sup>2</sup> Innogen, School of Social and Political Sciences, University of Edinburgh <sup>3</sup> Institute of Evolutionary Biology, School of Biological Sciences, University of Edinburgh <sup>4</sup> Present institutional address: Computing Science and Mathematics, Faculty of Natural Sciences, University of Stirling 9 \*Corresponding author. *E-mail address*: Kairsty.Topp@sruc.ac.uk 10 11 Keywords: spring barley, regression model, disease resistance, disease pressure 12 Abstract This paper assesses potential for Integrated Pest Management (IPM) techniques to 13 14 reduce the need for fungicide use without negatively impacting yields. The impacts of three 15 disease management practices of relevance to broad acre crops -disease resistance, 16 forecasting disease pressure, and fungicide use - were analysed to determine impact on 17 yield using a long-term field trials database of Scottish spring barley, with information from 18 experiments across the country regarding yield, disease levels, and fungicide treatment. Due to changes in data collection practices, data from 1996 – 2010 were only available at trial 19 20 level, while data from 2011 – 2014 were available at plot level. For this reason, data from 21 1996 – 2014 were analysed using regression models, while a subset of farmer relevant 22 varieties was taken from the 2011- 2014 data, and analysed using ANOVA, to provide additional information of particular relevance to current farm practice. While fungicide use 23 24 reduced disease severity in 51.4% of a farmer-relevant subset of trials run 2011 – 2014, and 25 yields were decreased by 0.62t/ha on average, this was not statistically significant in 65% of trials. Fungicide use had only a minor impact on profit in these trials, with an average 26 27 increase of 4.4% for malting and 4.7% for feed varieties, based on fungicide cost and yield 28 difference; potential savings such as reduced machinery costs were not considered, as these

29 may vary widely. Likewise, the1996 – 2014 database showed an average yield increase of 0.74t/ha due to fungicide use, across a wide range of years, sites, varieties, and climatic 30 31 conditions. A regression model was developed to assess key IPM and site factors which 32 influenced the difference between treated and untreated yields across this 18-year period. 33 Disease resistance, season rainfall, and combined disease severity of the three fungal 34 diseases were found to be significant factors in the model. Sowing only highly resistant varieties and, as technology improves, forecasting disease pressure based on anticipated 35 weather would help to reduce and optimise fungicide use. 36

37 I. Introduction

38 Fungicides are widely used in arable agriculture to reduce disease burden and its 39 impact on yields and quality, yet the effect of fungicides on yield is far from clear. While 40 some field studies show overall increases in yield (Kelley 2001 working on winter wheat; Paul et al. 2011, maize; Willyerd et al. 2015, winter wheat), others find no increase (Poysal, 41 42 Brammallz, and Pitblados 1993, tomato; Swoboda and Pedersen 2008, soybean), and many 43 present highly mixed results (Cook et al. 2002, wheat; Cook and King 1984, barley and wheat; Gaspar et al. 2014, soybean; Mycroft 1983, barley and wheat; Priestley and Bayles 44 1982, barley and wheat; Wiik 2009, winter wheat). Given that intensive fungicide use also 45 46 has a variety of concurrent detrimental effects, such as negative impacts on soil health and 47 soil ecosystems (Chen et al., 2001; Walia et al., 2014), and non-target toxicity linked to biodiversity loss in agricultural areas (McLaughlin & Mineau, 1995; Robinson & Sutherland, 48 49 2002; Geiger et al., 2010), alternative approaches to managing pests and diseases are 50 increasingly sought after. One such alternative is Integrated Pest Management (IPM), an 51 ecosystem based approach, first proposed by Stern et al. (1959), which combines diverse

management practices in order to minimize the use of pesticides while protecting crops from 52 53 pests and pathogens. IPM is an ecosystem approach which combines diverse management 54 practices in order to minimize the use of pesticides while protecting crops from pests and pathogens 55 (FAO 2017), and has been found to improve the overall environmental sustainability of farms, as 56 compared to conventional pesticide use situations (Lefebvre, Langrell, and Gomez-y-Paloma 2014). 57 IPM can encompass a number of techniques to reduce pathogen population levels or impact 58 on crops, including spraying pesticide where appropriate, crop rotation, varietal disease 59 resistance, forecasting disease pressure, adjusting product dose and timing, sowing 60 early/late in the season, and monitoring disease in field so that inputs can be adjusted accordingly. 61

In order to target and reduce fungicide inputs, while maintaining high yields, it is necessary to understand under what conditions (i.e. weather, varietal resistance level, previous crop, etc.) fungicide application impacts yields. Applications can then be tailored to situations where a yield increase is likely to occur, and eschewed when yield is unlikely to be impacted. An understanding of the situations in which various IPM strategies impact yields is also necessary, in order for uptake of these techniques to be optimised.

Proving direct links between fungicide use, yields, management strategies, and disease is difficult. For example, several experiments on wheat have linked fungicide use to yield increases. Work on fungicide control of powdery mildew (caused by *Blumeria graminis* f. sp. *tritici*) and septoria (caused by *Zymoseptoria tritici*) diseases found wheat yield increases of up to 2.7 t/ha (Jørgensen et al., 2000). Cook and King (1984) conducted field surveys of winter wheat, and found yield responses to fungicide use of up to 89%, with the most damaging leaf disease being mildew. However, many experiments have reported 75 inconsistent results - in wet conditions, for example, fungicide use increased yields in winter wheat grown in the US, while in dry years this was not seen (Wegulo et al., 2012). In a long-76 77 term field experiment on wheat in Sweden, only 52% of the years between 1983 and 2007 78 showed significant increases in yield from fungicide use (Wiik & Rosenqvist, 2010). Priestley 79 and Bayles (1982), working on spring barley in England found that yield impact from 80 fungicide use varied between years from a 2.4% increase in yield to 13.8%. The relationship between fungicide use, reduced disease, and increased yields therefore remains unclear, 81 82 complicating management decisions. A number of factors likely contribute to this variation, 83 including disease development, changes in yield potential, disease tolerance in the crop 84 (Bingham et al. 2009), and the physiological effects of fungicide on barley, which may be beneficial even in the absence of disease (Bingham et al. 2012). 85 86 Analysing data collected across a range of sites, in different fields, with different 87 weather conditions, and different management practices, can offer useful insight into which factors are most influential in determining the impact of treatment on yield. Much of the 88 literature on the use of key IPM techniques is based on experiments running for less than 89 90 five years (e.g. Makowski et al. 2005 working on sclerotinia in French oilseed rape; (Loyce et 91 al. 2008) working on diseases of French winter wheat; (Mazzilli et al. 2016) working on wheat in Uruguay). The work by Twengström et al. (1998) and Yuen et al. (1996) on 92 93 sclerotinia stem rot of oilseed rape is an example of an attempt to link yield and disease, 94 providing both a forecast of the likely disease severity and a risk algorithm, and considering 95 a range of factors, including crop rotation, rainfall, and previous disease incidence. Here, each factor was assessed first in an individual regression, then a full model was compiled, 96 97 including all terms, and a given factor removed to determine whether or not its inclusion improved the model's ability to predict epidemics (Twenström et al. 1998). While this work 98

provided a useful tool for farmer decision making, one issue which was specifically raised 99 by Twengström was the lack of data going back further than six years – longer term 100 101 experimental work was suggested as a way of improving predictive power. While few 102 studies on long-term data have thus far been conducted which explicitly test the impact of 103 fungicide use on yield and disease levels, Wiik and Ewaldz's (2009) work on winter wheat in Sweden using data from 1983 – 2005, followed by further analysis done by Wiik (2010) of the 104 data for 1977 – 2005 are notable exceptions, and both suggest that yield increases from 105 106 fungicide treatments are highly variable. Maximum yield increase from a single fungicide 107 treatment in 1983 – 2007 was found to be 1.9 t/ha and minimum yield increase was under 0.3 108 t/ha (Wiik & Ewaldz, 2009). Similarly, Cook and Thomas (1990), working on winter wheat in the UK, saw large fluctuations in yield response to fungicide across years, with one 109 110 fungicide application per season leading to average yield increases of 0.77 t/ha in 1985, but as little as 0.38 t/ha in 1984. Due to this variability, calls have been made for further analysis 111 of long-term field trials which compare yield, disease, and treatment, to allow optimisation 112 of fungicide use (Wiik, 2009). 113

Long-term databases can potentially provide useful information regarding IPM 114 115 efficacy, as data can be collected in a number of weather and agronomic situations, within the same region. However, assessing long-term data can be problematic, as data collection 116 117 and storage methods are likely to have changed over time, especially where the data has 118 been initially collected for purposes other than long-term analysis. In addition, the 119 institutional funding and dedication required to produce long-term datasets is often lacking, 120 due to other institutional pressures. Long-term datasets therefore often provide information 121 with varying levels of quality and consistency (Clutton-Brock and Sheldon 2010). Despite these drawbacks, the use of long-term data continues to be considered a useful way of 122

teasing apart complex relationships and causality in ecological studies (Clutton-Brock and
Sheldon 2010; Lindenmayer et al. 2012), and, along with information regarding betweenyear weather variation, can therefore provide a useful starting point for considering disease
prevention.

127 The present study makes use of a long-term field trials database collected regarding spring barley in Scotland to assess the impact of spraying fungicide and implementing IPM 128 on crop yields. Barley is one of the most widely grown crops in the world, with an average 129 of 53,572,792 hectares harvested each year, globally (FAOSTAT, 2013), and is of particular 130 importance in Scotland, where spring barley is the main cereal crop, accounting for 131 approximately 50% of arable land (excluding permanent grassland) in 2016 (Scottish 132 Government, 2016b). The key pests of barley are fungal pathogens, which have been estimated to 133 134 cause a total yield loss of 15% worldwide (Oerke and Dehne 2004) and 14% in the USA (James, Teng, 135 and Nutter 1991). To combat these diseases, a total of 187,173 kg of fungicide was applied to Scottish spring barley in 2014 representing 42% of the total amount of pesticide applied to the crop 136 (Scottish Government 2014). Fungicide use in Scottish spring barley therefore provides a useful case 137 138 study opportunity to assess the potential for reducing pesticide use, in a system which is of both 139 local and global importance.

Three fungal diseases of particular importance to spring barley production were
assessed as part of this work: mildew (caused by *Blumeria graminis* formae specialis *hordei*),
Rhynchosporium (caused by *Rhynchosporium commune*) and Ramularia (caused by *Ramularia collo-cygni*). Humidity has been proposed as a key risk factor for all three diseases (mildew:
Channon, 1981; Rhynchosporium: Ryan & Clare, 1975, Salamati & Magnus, 1997 ;
Ramularia: Havis et al, 2012 ), as have temperatures between 15 and 21°C (mildew: Polley

and King, 1973; Rhynchosporium: Salamati & Magnus 1997, Ryan & Clare, 1975, Xue & Hall,
1992; Ramularia: Havis et al., 2015).

148	Reducing fungicide use – if this can be achieved without impacting yields – could offer
149	an opportunity to reduce the negative environmental and health impacts associated with
150	crop production. This study aims to identify key management and environmental factors
151	which drive yield difference between sprayed and unsprayed spring barley. A basic
152	economic analysis is also presented to assess the potential impact on farmer's profits, had
153	they opted not to use fungicides in 2011 – 2014, providing insight into what is likely to be a
154	key driver of farmer behaviour.

- 155 II. Materials and methods
- 156

### a) Field Trials data as a platform for analysis

Data has been collected from field trials at a range of locations across Scotland since 1983 157 regarding yield, disease levels and fungicide treatment, along with a range of other 158 management factors. As the trials included widely used cultivars across this period, the 159 160 Field Trials database can provide a particularly farmer-relevant set of analyses. After an 161 extensive review of the Field Trials database, information from 1996 (the year in which 162 reports began to be stored electronically) onwards was retrieved for analysis; due to quality 163 issues in the older data, this paper analyses solely the information from 1996 - 2014 (see 164 Table 1 for a summary of the geographical spread of this database). Trials used a randomised block design with three or four replicates per trial and plots 165 ranging in size from 20 to 40m<sup>2</sup>. For each block within the trial, data for one untreated plot 166 was recorded in the database, alongside one fungicide treated: the 'best practice' treatment 167 168 for that year as determined by expert opinion (obtained from the lead plant pathologist at

169	Scotland's Rural College [SRUC], based on the results from the larger trials programme from
170	which this data set is extracted), allowing direct comparison of within-block differences
171	between treated and untreated plots. The 'best practice' treatment varied in chemistry,
172	timing, and number of applications between years and locations across the database. For
173	each trial in the Field Trials database, information is recorded about key farm management
174	features (e.g. varietal selection, preceding crop, sowing date, etc.), fungicide use information
175	(type, dose, and timing of application), disease information (percentage disease severity for
176	a number of key diseases at several growth stages during the crop growing season), and
177	yield. The number of disease assessments and the growth stages at which these were
178	measured during the growing season varied between trials, and by year and location. Trials
179	were assessed for disease at each application timing and usually 2-3 weekly thereafter until
180	the crop was senesced (less than 50% green leaf area on last remaining leaf). Though data
181	regarding the quality of the barley yield was collected for some trials, this was not
182	consistently recorded throughout the database, and so is not considered in these analyses.

Table 1: Summary of the geographical spread across Scottish Government sub-regions in
the 1996 – 2014 database

	Clyde Valley	Dumfries & Galloway	Fife	Lothian	North East	Scottish Borders	Tayside	Total trials in this year
1996				4		3		7
1997						1		1
1998				7				7
1999		1		2		2		5
2000				3		1	1	5
2001						1	1	2
2002				1			1	2
2003		2		1		1	1	5
2004		3	2	4			2	11
2005			1		1		1	3
2006						3	1	4
2007				2		3	1	6
2008						1		1

	Clyde Valley	Dumfries & Galloway	Fife	Lothian	North East	Scottish Borders	Tayside	Total trials in this year
2009							3	3
2010				2			1	3
2011	1		1	4			3	9
2012	2		1	6			1	10
2013	4			9			1	14
2014	5			7	1		1	14

185

186 b) Data collection and preparation

Additional data regarding weather, varietal disease resistance, and area under the 187 disease progress curve were added to the Field Trials database for analysis as described 188 below. Monthly regional weather data for each year were downloaded from the Met Office 189 190 for the two regions relevant to the trials database; Eastern and Western Scotland (Met Office, 191 2016). A list of the trial locations in each region is presented in Table 2. As anomaly weather data, showing variation from the mean, were not directly available from the Met office for 192 the growing seasons (March – August, inclusive, based on average growing season within 193 the Field Trials database), mean temperature and rainfall were calculated using Met Office 194 195 weather data for each region from 1981 – 2010, the most recent baseline available from the 196 Met Office, for the full growing season. Anomaly values were then calculated in accordance with the levels used in the Met Office (2016b) 1981 – 2010 anomaly maps (for more details on 197 198 the methods used to produce these maps, see Met Office 2016b). A growing season was therefore classed as 'wet' if the percent of average rainfall in that period was 110% or more, 199 and 'dry' if under 90% of the average; it was classed as 'hot' if more than 0.5°C higher than 200 average, and 'cold' if more than 0.5°C colder than average, as per the Met Office anomaly 201 202 map classes (see Table 3). Additional classifications of 'very hot' and 'very dry', etc. were

- trialled in initial stages of exploratory data analysis, but due to a lack of variability in the
- 204 weather, these were not used in the final version of the database.
- 205

### 206 Table 2: Regions corresponding to trial locations in the 2011 – 2014 database

Region	Trial location	Latitude	Longitude	Average yield (t/ha)	Average sow date
East of	Burnside BDE	56°28′ 56.40″ N	003°27′ 28.99″ W	5.8	75
Scotland	Balruddery BRY	56°28′ 55.77″ N	003°07′ 48.16″ W	7.1	82
	Balgonie BIE	56°11′ 02.65″ N	003°06′ 24.36″ W	6.5	83
	Boghall BLL	55°52′ 16.78″ N	003°12′ 29.25″ W	6.6	87
	Cauldshiel CEL	55°53′ 35.87″ N	002°50′ 04.68″ W	5.6	76
West of Scotland	Drumalbin DIN	55°37′ 26.80″ N	003°44′ 25.73″ W	6.9	93

207

# 208 Table 3: Rainfall and temperature anomalies for each region in the 2011 – 2014 database

Region	Growing season rainfall anomaly value 2011 2012 2013 2014			Grow 2011	ing season anomaly 2012	temperatur value 2013	ce 2014	
East of Scotland	Wet	Wet	Dry	Wet	Average	Cold	Average	Hot
West of Scotland	Wet	Wet	Dry	Average	Average	Average	Average	Hot

209

210 Varietal disease resistance information was added to the database using the

211 SRUC/Scottish Agricultural College & Home Grown Cereals Authority cereal recommended

- 212 lists for Scotland (1996 2014). Where a variety was not included in the recommended lists,
- and therefore could not be compared with other trials, it was removed from the database.

In order to provide a quantitative measure of disease intensity which could be used
to assess impact of fungicide use on disease, AUDPC was calculated using the standard
trapezoidal method, after Madden et al. (2007), such that:

217 
$$AUDPC = \sum_{j=1}^{n_j-1} \left(\frac{y_j + y_{j+1}}{2}\right) (t_{j+1} - t_j)$$

Where *tj* is the sample at a given time point *j*, *yj* is the disease level at the time point *j*, and *nj* is the number of time points. Growing season AUDPC was calculated for each of the three diseases (Rhynchosporium, Ramularia, and mildew) for each trial, as was Total AUDPC (the sum of AUDPC for the three diseases).

In a number of cases for trials prior to 2011, yield and disease severity measurements were recorded only as means for a given treatment, rather than at plot level. Where possible, plot level data was retrieved from old trial reports, but in a majority of cases plot level data was unavailable. A means database was therefore created, running from 1996 - 2014, by taking means of plot level data, where available, in order to render the database internally consistent.

Prior to analysis of the full dataset, a subset of the data chosen for its direct relevance to current commercial farmers was first analysed. This subset comprised the last four years of information available (2011 – 2014), for the varieties which were in use by farmers during this period (as determined by a farmer survey, reported in Stetkiewicz et al. (2018)) to provide information which is relevant to current farmer decision making. Data in this subset was available at individual plot level, which also allows for statistical analysis within trials, something which is not possible for the full dataset, due to the lack of plot level data.

#### 235

### c) Analysis of the 2011 – 2014 plot level subset

First, overall mean and median difference in yields between treated and untreated plots 236 237 in the Field Trials database were calculated using the within-trial block data, which was summarised for the variety. As an assessment of the impact of treatment on trial yields and 238 239 disease severity, ANOVA was conducted on each individual trial and variety combination, 240 using Genstat 16 (VSN International, 2013), and using within-trial block as the blocking 241 structure. The impact of treatment was tested for yield, mildew AUDPC, Ramularia AUDPC, Rhynchosporium AUDPC, and Total AUDPC. Significance was set at p<0.05. 242 A simple economic analysis was then conducted, using fungicide application cost data 243 244 (not including labour and machinery costs) from the SAC Farm Management Handbook calculations, which was available for spring barley in 2013 and 2014 (SAC Consulting, 2014; 245 246 SAC Consulting, 2013). For 2011 and 2012, fungicide cost data was not recorded separately 247 from total treatment costs, which included herbicides, insecticides, growth regulators and 248 trace elements (SAC Consulting, 2011; SAC Consulting, 2012). Fungicide applications represented, on average, 69.2% of the total application costs for the years 2013 – 2016 (SAC 249 250 Consulting, 2015; SAC Consulting, 2016; SAC Consulting, 2013; SAC Consulting, 2014). The 251 cost of fungicide applications in 2011 and 2012 was therefore assumed to be 69.2% of the total reported treatment costs. Spring barley price information was taken from the AHDB's 252 market data centre, where two-monthly average prices for spring barley were available 253 separately for both feed and malting varieties (AHDB, 2016c). Feed varieties were not 254 255 included in the Field Trials database for 2013 and 2014, meaning profit margin calculations 256 were not possible for this period. Average Scottish prices for each market type were 257 calculated by year for use in the analysis. This allowed a simple estimate of the difference in 258 profit per hectare between treated and untreated systems to be calculated. The impact of

fungicide treatment on difference in profit was assessed across the four years for eachvariety use type using two-way ANOVA.

# 261 d) Absolute yield difference regressions

262 Models

Stepwise regressions using GLM (generalised linear model) in Minitab 16 (2010) were elaborated for two databases: the full means Field Trials database (1996 – 2014), and the plot level Field Trials database (2011 – 2014). One of the objectives of this work was to compare which variables were included in the final stepwise regression for each of these datasets.

The 2011 – 2014 plot level data gave a high level of detail over a short period of time; this shortened period thus provided less factor variability to test, as there were necessarily a relatively small number of varieties, preceding crops, and weather conditions. Using the full dataset for 1996 – 2014 provided the opportunity to compare a larger number of factor levels, though with means rather than plot level data, and thus is useful for assessing a

273 wider range of potential management situations.

274 The regression model results presented in this paper are based on the yield 275 difference between treated and untreated plots/trials. For the 2011 - 2014 plot level data, this yield difference was calculated in order to compare within-block treated and untreated 276 277 yields; for the 1996 – 2014 means database, data were not available for within-block comparisons, and so yield differences are analysed at trial level (each trial was comprised of 278 one variety of spring barley). For a summary of the data types and analysis, see Table 4. 279 280 The variables included in the stepwise regressions were: sowing date; preceding crop - barley or non-barley; any resistance - disease resistance rating of seven or more to at least 281 one of the three diseases; AUDPC; and season rainfall and temperature anomaly levels of 282

283 wet/dry/average and hot/cold/average, respectively. A normal error distribution and

identity link function were used, as residuals were distributed relatively normally, as

285 determined by a review of standardized residual histograms and half-normal plots. Errors

286 likely to arise due to aliasing were identified, and these interactions were excluded from the

analysis. Random effects were unable to be fitted in the model.

While models were developed to consider the three individual diseases, in a majority of instances, a lack of data for mildew AUDPC through incomplete field recording meant trials without this information were removed from the analysis, rendering the results from these regressions misleading. As such, the results presented in this paper represent only those models which assessed Total AUDPC, rather than individual disease AUDPC.

293 Table 4: Summary of data types and analysis for each dataset

	1996 - 2014 dataset	2011 - 2014 dataset
Data available at	trial level	plot level
Data for	all varieties trialled in this	only farmer-relevant
	period	varieties
Analysis	stepwise regression	stepwise regression
		within-trial ANOVA

### 294 III. Results

295

### a. 2011 – 2014 plot level initial analysis

### 296 Fungicide treatment does not significantly impact yield in the majority of trials

297 Though treated plots had, on average, higher yields than untreated by 0.62 t/ha (see

Table 5), the majority of trials (65%) did not show a statistically significant impact of

- 299 fungicide treatment on yields. In cases where disease was present, disease severity,
- 300 particularly Total AUDPC, was more likely than yield to be reduced by the fungicide
- 301 treatment (see Table 6, below). The significance of treatment impact on yield varied across

- 302 years and locations, with 2013 (the only one of the four years with a growing season classed
- 303 as 'dry' in both East and West Scotland) having no trials showing a significant impact. Not

304 all diseases were present in every trial; the majority of instances where disease was not

- 305 recorded occurred in trials where treatment did not significantly impact yields.
- 306

# Table 5: Mean and median of the treated and untreated yields and the difference between treated and untreated yields of spring barley

	Mean yield (t/ha)	Standard error of mean (t/ha)	Median yield (t/ha)
Untreated	6.23	0.11	6.38
Treated	6.84	0.12	6.82
Difference	0.62		0.44

# 309

# 310 Table 6: Significance of impact of fungicide treatment on yield and disease severity\*

	Number of trials significantly different	Number of trials not significantly different	Percent of trials significantly different**	Number of trials with no disease pressure
Yield	14	26	35.0	
Total AUDPC (all	19	18	51.4	3
diseases)				
Rhynchosporium	17	19	47.2	4
AUDPC				
Ramularia AUDPC	13	13	50.0	14
Mildew AUDPC	6	11	35.3	23

311 \*Significance at p<0.05

\*\*Trials with no disease pressure (a value of zero) are not included in percentage

313 significantly different, nor in the number of trials (not) significantly different

### 314 Fungicide use increases profit only marginally

- 315 The simple economic analysis conducted compares the mean reduction in yields
- from a lack of use of fungicide to the cost saved by not purchasing fungicides, and assumes
- 317 barley quality for treated and untreated is the same. The resulting difference in profit
- between treated and untreated fields is small, averaging 4.4% (£50.30/ha) for malting

- 319 varieties and 4.7% (£56.80/ha) for feed varieties (see Table 7). Fungicide cost margins do vary
- 320 by year, with malting varieties having, for example, net losses in 2013, compared with the
- 321 +7.5% difference in profit in 2012. This difference in margin was significant at  $p \le 0.05$  for
- 322 distilling varieties, but was not significant for feed varieties (see Table 7). This analysis
- 323 disregards other possible savings from lack of treatment (e.g. lower labour costs).

# Table 7: Cost benefit analysis for malting and feed barley from 2011 – 2014 in Scotland, based on Field Trial database yields

	Mean Malting Barley Price (£/t)	Mean Feed Barley Price (£/t)	Difference in fungio malting	cide cost margin for varieties	Difference in fungicio vari	le cost margin for feed eties
			£/ha	%*	£/ha	%*
2011	193.1	152.1	83.7	6.1	102.4	8.1
2012	200.1	169.4	79.8	7.6	11.1	1.2
2013	145.4	140.2	-24.4	-2.8	-	-
2014	119.3	115.1	62.0	6.9	-	-
Overall	164.5	144.2	50.3ª	4.4	56.8	4.7

325 \*Percent difference is based on the treated profits<sup>a</sup> Indicates the relevant difference in cost margin is significant at  $p \le 0.05$ 

# **b. Modelling the full 1996 - 2014 dataset**

327 Yield Difference

The mean yield difference between treated and untreated across all trials in the 1996

**329** – 2014 dataset was 0.74 t/ha (standard error: 0.06).

### 330 Factors retained in the model for the full 1996 – 2014 dataset

- 331 Stepwise regressions developed for the 1996 2014 data identified Any Resistance,
- season rainfall, and disease severity as significant factors (see Table 8). Season rainfall had
- the highest  $R^2$  when tested individually (12.5%) and when removed from the model (5.7%).

Any Resistance had the second highest impact on R<sup>2</sup> (9.5% and 5.5%, respectively), and Total

AUDPC, the only other factor included in the model, had the third largest impact (5.2% and

**336** 4.3%, respectively).

# Table 8: Comparison of R<sup>2</sup> impact of significant factors in the 1996 – 2014 stepwise regressions and individual factor analyses

	Change in R <sup>2</sup> when removed from	R <sup>2</sup> when tested individually (%)
	the stepwise model (%)	
Any Resistance	5.5	9.5
Season rainfall	5.7	12.5
Total AUDPC	4.3	5.2

339

### 340 Regression models - comparisons

341	The final stepwise models for both the 1996 – 2014 means dataset and 2011 – 2014
342	plot level dataset included Total AUDPC, though other factors varied between the models
343	(see Table 9). Only the 1996 – 2014 dataset included Any Resistance, for example, while
344	growing season temperature was significant in only the 2011 – 2014 plot level data. For the
345	1996 – 2014 means dataset there was complete agreement between the stepwise models and

- the individual factor regressions. The 2011 2014 plot level dataset had only one factor
- 347 which was significant when tested individually, but which did not remain in the stepwise
- 348 model: growing season rainfall.

349	Table 9: Final stepwise regressions	for each dataset,	including Total AUDPC*
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Model 1 – stepwise regression (1996 – 2014) including
Total AUDPC

g Model 2 – stepwise regression (2011 – 2014 plot level data) including Total AUDPC

	Significance	Coefficient	Difference to R <sup>2</sup> when removed from model (%)	Significance	Coefficient	Difference to R <sup>2</sup> when removed from model (%)
Season rainfall	Wet: 0.017 Dry: 0.110	0.2187 -0.186	-5.7			
Season temperature				Hot: 0.009 Cold: N/A	0.291	-3.8
Any Resistance	<0.001	-0.2817	-5.5			
Total AUDPC	<0.001	0.000489	-4.3	< 0.001	0.000574	-13.4
Model R <sup>2</sup>	21.2%			22.3%		

\*Factors highlighted in solid grey were significant in both the stepwise regression model and the individual regressions. Those with grey
dots as highlights were significant only individually. Significance was tested at p<0.05.</li>

352 IV. Discussion

353

### a) Fungicide treatment impact on yield is variable

The mean impact of fungicide treatment on yields from 2011 -2014 was 0.62 t/ha, 354 however, the difference in yield between treated and untreated was statistically significant 355 only 35% of the time. From 1996 – 2014, mean yield difference was 0.74 t/ha. Farmer survey 356 357 work indicates that most Scottish spring barley farmers estimated the yield benefit from 358 applying fungicides to be between 1 and 2 t/ha (Stetkiewicz et al. 2018), suggesting that if this yield difference is representative, farmers are overestimating the effect of fungicide. 359 Preliminary economic analysis suggests that increased profit from sprayed fields is in 360 361 the range of 4.5% for malting barley, considering only the difference between mean treated and untreated yields, and the cost of applying fungicides. When additional factors, such as 362 363 labour and machinery costs are taken into account, this figure may decrease further. This 364 analysis assumes that all untreated barley in the Field Trials was of sufficient quality for 365 malting, which may be inaccurate. There are, however, instances where fungicide treated yields were substantially (up to 2.01 t/ha) greater than those for untreated plots. In these 366 situations, for example where varietal disease resistance scores are low, or in years of 367 368 particularly wet weather, the scope for fungicide reduction or elimination is likely limited. Similarly, Wiik and Rosenqvist (2010) found that mean net return from fungicide use on 369 winter wheat in Sweden was 12 euro/ha over the 25 years studied, with mean net return 370 371 being negative in 10 years and with fewer than half of trials in 11 years being profitable to treat. Recent work on winter wheat in Sweden found that rain, disease severity, soil type 372 and previous crop were able to identify situations where fungicide treatment gave a positive 373 374 marginal return, and that profitability varied with wheat prices (Djurle, Twengström, and

Andersson 2018). Additional information about the costs, risks, and potential benefits would 375 give farmers more confidence when deciding whether or not to reduce fungicide inputs. 376 377 Approximately half of the 2011 – 2014 trials showed a significant impact of fungicide 378 treatment on Rhynchosporium, Ramularia, mildew, and Total AUDPC levels. Fungicide 379 treatment therefore appears to impact disease severity in a large number of trials, but this impact does not translate directly into a significant impact on yield. Disease tolerance, 380 whereby the yield of some genotypes is less affected by a given level of disease than other 381 genotypes (Bingham et al. 2009 working on barley and wheat), may explain some of this 382 variation. Treatment significance varied across year and location, suggesting other factors 383 also impact yield difference, such as, perhaps, soil type and quality. Further, 2013, the driest 384 year, and therefore a year which was not conducive to fungal growth, was also the only year 385 with no trials showing a significant impact of treatment on yield. Previous work on long-386 term databases of winter wheat has found precipitation, along with temperature, to be a 387 significant factor in predicting yield and disease severity (Wiik & Ewaldz, 2009). 388

389

### b) Key factors influencing impact of fungicides on yield

The results from the 1996 – 2014 regression model suggest that using season rainfall (perhaps via a model using within-season weather to identify periods of high risk, as done for Sclerotinia stem rot in oil seed rape by Yuen et al. (1996), a project which falls beyond the scope of this paper) as an indicator for likely need to spray fungicide, in conjunction with varietal disease resistance, has the potential to reduce the need for fungicide use while maintaining high yields. In all stepwise and individual factor regression models, regardless of the dataset tested, Total AUDPC was identified as an important factor in terms of yield difference between treated and untreated trials, suggesting that where fungicide use iseffective at increasing yields, this may be related to its reduction of disease severity.

High levels of resistance to one or more of the three diseases was also important in both
stepwise and individual factor regression models developed for the full 1996 – 2014 dataset.
In all cases disease resistance was linked with lower yield differences between treated and
untreated trials. That disease resistance buffers the effect of not spraying fungicide is well
established in the field trial literature for wheat diseases (Berry et al., 2008; Cook & Thomas,
1990; Martens et al., 2014).

405 Season rainfall remained in the Total AUDPC stepwise regression model developed for 406 the 1996 – 2014 means level data, where wet seasons were linked with larger yield differences between treated and untreated. Similarly, dry seasons were linked with smaller 407 408 yield differences between treated and untreated in the plot level individual regressions. Dry 409 conditions have previously been seen to lower the impact of fungicide use on wheat yields in long-term experiments (Wiik & Ewaldz, 2009), and to be crucial to high yields in Scottish 410 barley (Brown, 2013), while wet periods have been proposed as one of the risk factors for 411 412 Ramularia (Havis et al., 2015) and Rhynchosporium (Ryan & Clare, 1975; Xue & Hall, 1992) 413 to flourish, as has humidity for mildew development (Channon, 1981), conclusions which are supported by this analysis. 414

Final stepwise regression models were related to individual factor regressions, following a similar method used to assess risk factors for sclerotinia in oilseed rape using logistic regressions (Yuen et al., 1996). For both datasets, the Total AUDPC stepwise regressions fitted the individual factor regression results well, with five out of the six factors which were significant when tested individually also being retained in the relevant stepwise model (those retained in the 1996 – 2014 dataset analysis: growing season rainfall, Any Resistance and Total AUDPC; those retained in the 2011 – 2014 dataset analysis: growing season
temperature, Total AUDPC; season rainfall was significant individually for the 2011-2014
data, but not retained in the model).

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### c) Parallels and differences in results from the two datasets

425 The final stepwise models for both datasets using Total AUDPC were similar: each 426 included Total AUDPC and one weather variable (season temperature for the 2011 - 2014 plot level data, and season rainfall for the full 1996 – 2014 dataset), though Any Resistance 427 428 was only included in the full 1996 – 2014 dataset model. As the only stepwise model for 429 Total AUDPC which contained a factor not significant when tested in an individual 430 regression (season temperature) was that created for the 2011 - 2014 plot level data, it is not 431 clear that plot level information provides a more accurate representation of the factors influencing yield difference than mean, trial-level information. In this instance, means level 432 433 long-term data seems to provide more useful results for understanding the impact of management and weather factors on yield differences, due to the larger amounts of 434 variation than are seen in the short term database. In future, comparing results from a long-435 436 term plot level database and its means counterpart could provide useful data about which is more important in modelling factor impacts on yield. 437

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### d) Limitations

A number of limitations to this study exist which are, in large part, due to the difficulties inherent in using a large database which has been collected for other purposes. Few conclusions can be drawn from this work regarding the potential influence of sowing date and preceding crop on disease and yield impacts of fungicide application, due to a lack of variation in the database for these factors. An attempt was made to include early season

disease measurements (between GS 24 - 34) as a way of considering disease which provides 444 farmers with a measure to act upon within season, as recommended in previous decision 445 446 making tools (Burke & Dunne, 2008), however a lack of sufficient data prevented this from 447 inclusion in the regressions analysis. More information regarding these factors, as well as 448 more detailed weather data, linked to each individual farm or county, rather than data compiled at regional level, could provide more insight into the factors of interest. 449 In addition, the small size of plots included in the Field Trials database (typically 20 x 450 451 2m), as compared to the size of a commercial barley field, combined with the fact that the single untreated plot in any given trial block is surrounded by treated plots, may reduce the 452 453 yield difference between treated and untreated plots by buffering the plot from disease pressure. Within the models themselves, being unable to include random terms, or 454 interactions between terms such as rainfall and temperature (which are unlikely to be fully 455 independent) also restricts the robustness of the results. Assessing diseases at an individual, 456 rather than aggregate level could also provide more precise results, which may be of value 457 in management decisions. 458

The use of large datasets such as the Field Trials database provides opportunities for analysing variation across a wider range of conditions, but, as many of these long-term data sources were not designed with such analysis in mind, the lack of potentially useful detail is an important trade-off of using such data. Despite these limitations, and though finer detail could no doubt be revealed with additional data, important patterns regarding the impact of fungicide use on yield were detected. 465 V. Conclusion

Fungicide treatment impacted yield levels significantly in just over one third of the trials 466 467 assessed from 2011 – 2014, though disease levels were significantly reduced in many cases. 468 The lack of a constant influence on yield, and the minimal cost benefit from fungicide 469 treatment, estimated at less than 5% on average, suggests there may be an opportunity to 470 reduce fungicide use in this sector with little negative impact on yield or profit. In addition, the yield differences seen in these field trials (on average: 0.62 t/ha for 471 commercially relevant varieties grown from 2011 – 2014 and 0.74 t/ha for all trials in the 1996 472 473 – 2014 database) were well below those expected by Scottish spring barley farmers and 474 agronomists (Stetkiewicz et al. 2018). Stetkiewicz et al. (2018) report 71.8% of surveyed 475 farmers and 75% of agronomists estimating the impact of fungicide application to spring 476 barley to be between 1 and 2 tonnes per hectare – well above the impacts reported here. 477 Farmers and agronomists therefore appear to be substantially overestimating the impact of 478 fungicide use on yield. 479 Using the final stepwise regression model developed for the full 1996 – 2014 dataset 480 testing Total AUDPC, and the individual regressions for this data, three factors appear to be 481 crucial in determining the impact of fungicide treatment on yield in the Field Trials database: season rainfall, disease resistance, and Total AUDPC. Ranked by R<sup>2</sup>, season 482 rainfall explains the most variation in yield difference, followed by Any Resistance, and 483

Total AUDPC. As fungicide use did not always result in increased yield, and the increases
which did occur were often minimal, forecasting disease severity for the season and acting
upon this, e.g. planning to spray when the season is forecast to be wet and reducing
spraying when dry, may help to rationalise fungicide use, given that the alternative of
waiting until a disease appears before treating would preclude the use of preventative

489	fungicides, and restrict available products to those with curative action. Similarly, sowing
490	only spring barley varieties which are highly resistant to one or more key diseases may
491	reduce the need for fungicides. The inclusion of Total AUDPC as a key factor highlights the
492	fact that disease severity is important in yield dynamics; this may be managed within season
493	through a combination of techniques, including fungicide applications. Other IPM measures,
494	such as rotation and sowing date, may play a role in determining yield impacts of
495	fungicides, but could not be fully assessed here, due to lack of variation. These models
496	provide a useful tool for assessing the relative merits of different IPM techniques on yield
497	and allow farmers and decision makers to prioritise acting on those which have a significant
498	explanatory effect, such as sowing highly disease resistant varieties.

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