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All-cause and cause-specific mortality among people with regular or problematic cocaine use: A systematic review and meta-analysis

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Cocaine use and mortality

All-cause and cause-specific mortality among people with regular or problematic cocaine use: A

systematic review and meta-analysis

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ABSTRACT

Aims: To estimate pooled all-cause and cause-specific mortality risk for people with regular or problematic cocaine use.

Methods: Systematic review and meta-analysis of prospective or retrospective cohort studies (n≥30) of people with regular or problematic cocaine use with data on all-cause or cause-specific mortality. Of 2808 papers, 28 were eligible and reported on 21 cohorts with a total 170,019 individuals. Cohorts identified based on acute care for drug poisoning or other severe health presentation were excluded. Title/abstract screening was conducted by one reviewer; a second reviewer independently checked 10% of excluded studies. Two reviewers conducted full-text screening. Data were extracted by one reviewer and checked by a second. A customised review-specific study reporting quality/risk of bias tool was used. Data on crude mortality rates (CMR) and standardised mortality ratios were extracted for both all-cause and cause-specific mortality. Standardised mortality ratios were imputed where not provided by the author using extracted data and information from the Global Burden of Disease Study 2017. Data were pooled using a random-effects model.

Results: The pooled all-cause crude mortality rate was 1.24 per 100 person-years (95% CI: 0.86, 1.78; n=16 cohorts), but with considerable heterogeneity (I²=98.8%). The pooled all-cause standardised mortality ratio was 6.13 (95%CI: 4.15, 9.05; n=16 cohorts). Suicide (SMR 6.26, 95%CI 2.84, 9.68), accidental injury (SMR 6.36, 95%CI 4.18, 9.68), homicide (9.38, 95%CI 9.38, 3.45, 25.48), and AIDS-related mortality (SMR 23.12, 95%CI 11.30, 47.31) were all elevated compared with age and sex peers in the general population.

Conclusions: There are elevated rates of mortality among people with regular or problematic cocaine use for traumatic deaths and deaths attributable to infectious disease.

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INTRODUCTION

Cocaine manufacture is estimated to be at the highest level recorded and trafficking routes continue to expand globally (1). Greater availability has been accompanied by increased use, with an estimated 0.37% (CI 0.31-0.42) of adults aged 15-64 years reporting use globally in 2017, equivalent to 18.0 (CI 15.5-21.0) million people (1). Confluence of increasing availability and demand, coupled with indicators of increasing purity in some regions (1, 2), creates a higher risk environment for increased problematic use and health harms.

The most common forms of cocaine are hydrochloride salt (typically a fine white powder) and base ('crack' cocaine), the latter appearing crystal- or rock-like and typically being higher purity (3). Cocaine produces stimulant effects, increasing heart rate and blood pressure, enhancing alertness and producing feelings of euphoria, with a half-life of 30-60 minutes (4). Regular use of cocaine is associated with adverse health effects, predominantly cardiovascular (e.g., arrhythmia, myocardial infarction, stroke) and psychiatric (e.g., psychotic episodes, suicidal ideation) in nature (5, 6). There are additional and greater severity health risks depending on route of administration, including respiratory problems with smoking (6), nasal ulceration with snorting (7) and bloodborne virus transmission with injection, as well as via form used (e.g., higher risk of dependence with crack versus powder cocaine; 6).

A systematic review of studies published until 2008 suggested that people with regular cocaine use had a four to eight times higher mortality risk than their age and sex peers in the general population(8). Only seven cohorts were located, so a lack of data precluded quantifying risks of cause-specific mortality or examination of reasons for variation in mortality rates. In the intervening period, there has been significant research exploring specific harms associated with cocaine use from which causespecific mortality estimates could be computed. Increased cocaine supply and demand globally and increasing poisoning deaths in a number of countries (e.g., United States; 9), including the highest rate of deaths in England and Wales observed in 2017 (10), makes it imperative that the magnitude of

mortality risk by cause be quantified for this population. As such, the aims of this systematic review were to:

- 1. Update the previous systematic review and perform meta-analysis to calculate pooled estimates of all-cause crude mortality rates (CMRs) and standardised mortality ratios (SMRs) for people with regular or problematic cocaine use globally (regardless of treatment status); and
- 2. Compute pooled estimates of cause-specific CMRs and SMRs, with a focus on those causes of death that could be causally related to regular or problematic cocaine use.

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METHODS

We report the systematic review methodology in accordance with the PRISMA guidelines (11) (**Appendix A**). This review protocol was registered with PROSPERO (CRD42018094623; **Appendix B**).

Search strategy and study screening

Medline, Embase and PsycINFO peer-reviewed literature databases were searched using the OVID[™] interface/platform for articles published between 2009 and 22 February 2018. Relevant articles published between 1980 and 2008 were identified through a previous review conducted by the research team (8). In line with the previous review, search strings incorporating keywords and Medical Subject Headings (MeSH terms) related to cocaine/crack cocaine and mortality epidemiology were used to identify relevant articles (see **Appendix C and D**). Searches were limited to human literature. There were no restrictions on publication type or language; papers published in languages other than English were read via Google Translate or by a team member fluent in the language.

Citations were imported into an Endnote[™] library where duplicate citations were removed, and imported into Covidence, a web-based screening tool (12). Titles and abstracts were reviewed by one team member (LTT or research assistants); 10% of excluded studies were checked by a second person (AP) to monitor accuracy. Full-text articles were reviewed by two reviewers (LTT and AP or LD); discrepancies were resolved by a third reviewer (AP or LD). Reference lists for relevant systematic reviews identified in the peer-review literature search were hand searched for additional papers not already identified.

Study eligibility

Studies were included if they were cohort or case-control studies or clinical trials ($n \ge 30$ people) where: i) at least 90% of the sample reported regular or problematic use of cocaine and ii) data on all-cause or cause-specific CMRs or SMRs were available (see **Appendix E** for detailed inclusion and exclusion criteria). This could include cohorts that were identified based on criteria other than cocaine use but

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reported mortality data for a sub-group of people reporting regular or problematic cocaine use. Cohorts were defined as comprising people with regular or problematic use if individuals reported cocaine as their primary drug used, cocaine injection, cocaine dependence, treatment for cocaine dependence or other healthcare presentation related to the effects of cocaine. Cohorts defined by any cocaine use (with no other indicators of cocaine-related problems) were excluded. Cohorts defined by cocaine overdose/poisoning or other serious adverse health presentations with high mortality risk were excluded. This included cohorts defined by cardiovascular presentations related to cocaine exposure or HIV positive status; estimates from these studies are available in **Appendix F**.

Data extraction

The data extraction worksheet was developed in Microsoft Excel based on the previous review. extraction was standardised through the use of a manual detailing data entry rules (8). Data were independently extracted by one reviewer (LTT) and checked by a second reviewer (TS).

Variables extracted included study information (e.g., country of sample, length of follow-up, recruitment setting) and sample information (e.g., age, sex, percentage of sample injecting, percentage engaged in treatment, form of cocaine used). With respect to the outcomes of interest (all-cause and cause-specific mortality), we extracted sample size, number of observed deaths (all-cause and cause-specific), person-years of observation, CMRs and SMRs. Data were extracted for specific causes of death which might be causally related to regular or problematic cocaine use through direct acute or chronic effects of the drug, or indirectly via other risk pathways (e.g., use of non-sterile injection equipment, high risk sexual behaviours). These causes of death comprised: drug-related, accidental injury, suicide, homicide, cardiovascular disease, respiratory disease, cancer, AIDS-related and digestive disease (the latter including liver diseases)(6). We also extracted International Classification of Diseases (ICD) codes or other information used to define cause-specific mortality.

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to locate missing data. If supplementary information could not be located or did not detail the data needed, authors were contacted by email for additional information.

Quality of reporting and risk of bias

There are no standard tools available for assessing risk of bias in descriptive studies reporting estimates of CMRs or SMRs. However, risk of bias tools for observational studies assessing the effects of exposures are under development and some domains of these tools are clearly relevant (13). We assessed risk of bias for the two domains that we considered most relevant for this type of study design: sample representativeness (i.e., the number of site locations and sample types) and outcome measurement (i.e., ascertainment of death by death registry/certificate versus indirect sources). Studies were rated as being at higher or lower risk of bias on these domains. The tool also measured three components of study reporting quality, comprising cohort description (i.e., age and/or sex data for the cohort), data completeness (i.e., ICD codes or other information to define cause of death). Studies were assessed as having higher or lower risk or bias or quality reporting on each of these domains (see **Appendix G** for full details of tool).

Data analysis

Calculation of CMR, SMR, and relative risks

If not reported by study authors, CMRs were calculated as per 100 person-years (100PY). Where person-years were not reported nor made available by the authors, an imputed CMR follow up calculation was undertaken using other data reported by the authors. This calculation was based on the assumption that deaths occurred halfway through the follow-up period so that each person who died contributed half the person-year follow-up of survivors (see Appendix H for formula and example computation).

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Standardised mortality ratios (SMRs) represent the CMR ratio between those exposed to the risk and the general population (including those exposed to the risk). If not reported by study authors, SMRs were calculated by dividing the sample CMR by the CMR for the for the respective age, sex, location, and years from the Global Burden of Disease (GBD) 2017 study (14). The GBD 2017 study used vital registration, verbal autopsy, registry, survey, policy, and surveillance data to model mortality estimates for 282 causes of death in 195 countries and territories. Due to the nature of GBD data classification, a match between the GBD database and study classification for respiratory and digestive disease mortality was unable to be obtained resulting in CMRs being unable to be converted into SMRs. Standard errors of log-transformed CMRs and SMRs were estimated using Rothman Greenland method (15).

Relative risks (RRs) comprise the ratio of mortality risk between those exposed to the risk (i.e., regular or problematic cocaine use) and those not exposed to the risk. With a low prevalence exposure such as regular and/or problematic cocaine use, SMRs and RRs should be similar, but RRs provide an estimate that is useful for further estimation of burden of disease. RRs were estimated from SMRs using the method described by Jones and Swerdlow (16) by adjusting the SMR by the proportion of the general population that experiences cocaine dependence (see **Appendix H** for the formulae).

Pooling all-cause and cause-specific mortality CMR and SMR

DerSimonian and Laird random-effects meta-analysis (17) was conducted in STATA version 14.2(18) to pool all-cause and cause-specific CMR and SMR estimates. This allows for heterogeneity between and within studies (noting that each study could only contribute one estimate to each meta-analysis). Heterogeneity was quantified using the l² statistic and described as low (\leq 30%), moderate (>30% and \leq 50%), substantial (>50% and \leq 90%) or considerable (>75% and 100%)(19).

Understanding variation in CMR and SMR

Sources of heterogeneity in all-cause CMR and SMR were investigated through univariate metaregressions, including aspects of the study design (including the percentage of the cohort that was

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female, type of cocaine used, and geographical region) and the characteristics of the sample (including year of final follow-up, and recruitment setting). Variables were only included in meta-regressions if 5 or more data points were available per variable.

Sensitivity analyses

We undertook several addition analyses to examine potential impacts of our imputed metrics (CMRs, person years, SMRs) for studies that did not report all of these metrics upon the estimates generated in our meta-analyses. First, we generated estimated person years and CMRs for studies using our methods described above for studies that had reported all of these and compared the resulting estimates with the author-reported estimates; we found reasonable consistency for four out of seven studies (**Appendix H**). Second, we compared pooled estimated CMRs that included studies with imputed CMRs with the pooled estimates that only included author-reported CMRs. Finally, we contrasted the pooled estimated SMRs including studies where we used derived SMRs with pooled estimates that only used author-reported SMRs.

RESULTS

As evident in the PRISMA study flow diagram in **Figure 1**, the search identified 2,808 papers, of which 209 were screened in full. Of these 28 papers were deemed eligible, including 21 cohorts reported in primary publications and 7 secondary publications providing additional data for those cohorts reported in the primary publication (see **Appendix I** for excluded studies and **Table 1** for details of included studies).

All cohorts (or subsample of cohorts) included here wholly comprised people who reported regular or problematic cocaine use (see **Appendix J**). Cohort size ranged from 63 to 83,808 individuals (total 170,019 individuals) and follow up of the cohort ranged between 58 to 468560 PY (total 1,012,147 PY). Cohorts were recruited from nine countries: eight of which were classified as high-income and one (Brazil) as middle-high income. Over two-fifths of the cohorts (43%, n=9) were recruited from the United States, three cohorts from Spain and two from Denmark. Relative to the previous review (9 cohorts)(8), 13 new cohorts were identified and three cohorts had additional published follow-up data. One study(20) from the original review was excluded as, on review for the current study, the population was not identified as reporting regular or problematic cocaine use.

[Figure 1 here]

[Table 1 here]

Risk of bias and study reporting quality

Two-fifths (42.9%) of the cohorts were at risk of bias from poor representativeness; that is, the sample were recruited from a single site location or from a single sample type (e.g., people who were in treatment; see **Table 2**). Only three cohorts (14.9%) did not use an official death registry to identify deaths within the cohort, suggesting low risk of bias in outcome ascertainment.

Outcomes for quality of reporting were mixed. Both age and sex characteristics of the cohort were reported for all except three cohorts (14.9%). However, half (52.4%) of the cohorts had incomplete

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mortality data, where the numerator or denominator for the main analyses of mortality were not reported. Of the 17 cohorts for whom cause-specific mortality was reported, 10 (58.8%) did not provide definition(s) of the cause of death (see **Appendix J** for ICD codes used).

[Table 2 here]

It is important to note that quality of reporting outcomes was assessed for the total cohort. In many instances, people with regular or problematic cocaine use were a subsample of a larger cohort of people who use drugs. Age/sex characteristics and both numerator and denominator for mortality specifically for the sample who report regular or problematic cocaine use were thus only available for 8 (38%) and 5 (24%) cohorts, respectively (**Table 2**).

All-cause mortality

Data from 16 cohorts yielded an all-cause CMR of 1.24 per 100 PY (95% CI: 0.86, 1.78) with substantial heterogeneity (I²=98.8%) (**Table 3; Figure 2**). Meta-regression analyses to explore sample and study characteristics that could explain high heterogeneity showed that the percentage of the cohort reporting injecting drug use was positively associated with higher CMRs (**Table 4; Appendix L**). Similarly, recruitment of the cohort from a single city (versus subnational/national recruitment) was associated with higher CMRs. There was little or no statistical evidence of an association of CMR with other variables explored (**Table 4**).

[Figure 2 here]

Authors of studies for seven cohorts reported all-cause SMRs; we imputed all-cause SMRs for a further 9 cohorts. The pooled all-cause SMR across the 16 cohorts was 6.13 (95%CI: 4.15, 9.05), with substantial heterogeneity ($I^2 = 99.0\%$) (**Table 3, Figure 2**). This estimate was similar to that observed from the seven cohorts where all-cause SMRs were reported by the authors (5.58, 95%CI: 3,90, 7.99; $I^2 = 96.6\%$) (**Appendix L**). Excess mortality was particularly elevated among cohorts reporting lifetime injecting (13.74, 95% CI: 12.67, 14.90; $I^2=91.2\%$; n=5 cohorts).

[Table 3 here]

The percentage of the cohort that was female was negatively associated with excess mortality (**Table 4**; **Appendix L**). Study characteristics associated with excess mortality comprised recruitment setting (with higher SMRs for cohorts recruited from hospital versus treatment services and other settings) and sampling frame (with higher SMRs for cohorts recruited from cohorts recruited from a single city versus subnational/national recruitment).

[Table 4 here]

Drug-related deaths

Eight cohorts reported data on drug-related deaths. The definition of drug-related deaths was not provided for six cohorts; for the remaining two cohorts, 'drug-related deaths' comprised poisoning deaths, deaths attributed to mental and behavioural disorders due to psychoactive substance use, and deaths from other causes that were thought to be drug-related by the authors following consultation with forensic and toxicological services (e.g., ICD-10 code J81 'pulmonary oedema')(21). The pooled drug-related CMR was 0.34 per 100PY (95%CI: 0.10, 1.15), again with considerable heterogeneity (I² = 98.6%) (**Table 3, Figure 3**). SMRs were only reported for two cohorts, with a very high pooled estimate observed (44.37, 95% CI: 37.28, 52.81; I²=63.7%) (**Appendix L**). We have not reported the pooled estimate for all cohorts - including those for whom we imputed SMR as the estimate was deemed unstable.

[Figure 3 here]

Traumatic deaths: suicide, accidental injury, and homicide

The pooled suicide CMR was 0.07 per 100PY (95%CI: 0.04, 0.10; $I^2=72.6\%$; n=8 cohorts), with both pooled accidental injury (n=6 cohorts) and homicide (n=3 cohorts) CMR estimated at 0.09 per 100PY (95%CI: 0.04, 0.22; $I^2=95.4\%$ and 95%CI: 0.01, 1.54; $I^2=98.4\%$ respectively) (**Table 3, Figure 3**). The pooled SMR based on author-reported and imputed estimates was 6.26 (95% CI: 2.84, 13.80; $I^2=94.4\%$)

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for suicide, 6.36 (95% CI: 4.18, 9.68; I²=71.9%) for accidental injury and 9.38 (95% CI: 3.45, 25.48; I²=90.1%) for homicide (**Table 3, Figure 4**). Pooled estimates of the subsample of author-reported SMRs were higher for suicide and lower for accidental injury and homicide (**Appendix L**) but fell within the confidence intervals of the former estimates. There was substantial heterogeneity in these pooled CMR and SMR estimates.

AIDS-related deaths

The pooled CMR for AIDS-related deaths based on 6 cohorts was 0.28 per 100PY (0.12-0.63; $I^2 = 95.1\%$) (**Table 3, Figure 3**). Only one study reported a SMR for AIDS-related deaths, with an excess mortality rate five times that expected in the general population (4.98, 95% CI 0.70, 35.34; **Appendix L**). The pooled SMR based on author-reported and imputed estimates was higher (23.12, 95%CI 11.30-47.31; $I^2 = 90.1\%$; n=6 cohorts) (**Table 3, Figure 4**) but fell within the confidence interval for the former estimate.

Digestive disease deaths

The pooled CMR for digestive diseases based on three cohorts was 0.14 per 100PY (95%CI: 0.03, 0.73; $I^2 = 96.1\%$). The SMR was only reported by authors for one cohort (1.90, 95%CI 1.20, 2.89). Three different cohorts provided CMRs and SMRs for liver-specific deaths, with these deaths occurring more than three times the expected rate (pooled SMR: 3.36, 95%CI 0.51, 22.10; $I^2 = 92.5\%$; n=3 cohorts) (Table 3, Figure 4).

Other causes of death

Pooled CMRs were derived from four cohorts for cardiovascular disease (0.13 per 100PY; 95%CI: 0.07, 0.24; $I^2 = 77.3\%$), five cohorts for cancer (0.11, 95% CI: 0.50-0.25; $I^2 = 87.1\%$), and five cohorts for respiratory diseases (0.09, 95%CI 0.04, 0.17; $I^2 = 63.2\%$) (**Table 3; Figure 3**). These deaths were elevated relative to the expected rate, particularly respiratory diseases (pooled SMR: 24.12, 95% CI 6.03, 96.43) (**Table 3, Figure 4**), although this estimate should be treated with caution being based on one cohort.

[Figure 4]

Mortality relative risks

Estimated study-level mortality RRs for all-cause and cause-specific mortality are presented in **Appendix K** and are similar to SMRs.

We conducted several additional analyses to examine our approach to imputation of data. First, we generated estimated CMRs and person years for studies that had already reported all these metrics and found reasonable consistency across these (**Appendix H**). Second, we compared pooled estimates that did not include imputed CMRs with those that did, and again found reasonable consistency (**Appendix L**). Finally, as noted earlier, we also contrasted pooled SMRs that only included author-reported SMRs with those that included our imputed SMRs (**Appendix L**); in some cases there were very few or only one cohort that had author-reported SMRs, so there are few data, but the all-cause estimates were remarkably similar. Finally, we examined the potential impact of the two very large cohorts on pooled estimates by generating pooled estimates without those two cohorts (22, 23). Pooled all-cause CMR was 1.23 per 100PY (95%CI 0.75, 2.02; I²=98.7%; n=15 cohorts) compared to 1.24 per 100PY (95% CI: 0.86, 1.78; I²=98.8%; n=16 cohorts) including those cohorts; pooled accidental injury CMR was 0.13 per 100PY (95%CI 0.06, 0.25; I² = 82.5%; n=5 cohorts) compared to 0.09 per 100PY (95%CI: 0.04, 0.22; I²=95.4%; n=6 cohorts) including those cohorts; and pooled suicide CMR was 0.08 per 100PY (95%CI 0.04, 0.15; I²=73.0%' n=7 cohorts) compared to 0.07 per 100PY (95%CI: 0.04, 0.10; I²=72.6%; n=8 cohorts) including those cohorts.

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DISCUSSION

This review suggests that people with regular or problematic cocaine use have, on average, an excess mortality risk six times (95%CI four to nine times) the expected rate for their age-matched counterparts in the general population. The most highly elevated causes of deaths were drug-related or arising from traumatic causes (suicide, accidental injury or homicide) – all of which are preventable. Mortality was also elevated for communicable diseases (e.g., AIDS-related mortality) and other natural causes of death, including cardiovascular disease. These findings are consistent with the known effects of cocaine (6) and other risk pathways associated with these health outcomes (6). The potential for a growing population of people with regular or problematic cocaine use – coupled with the current findings of elevated mortality risk – reinforces the need for expansion of evidence-based prevention and intervention efforts to reduce health harms.

It is important to note that it is not the case that elevations in mortality necessarily reflect direct causal effects of cocaine use. In some instances elevated mortality may reflect other lifestyle factors and exposure to risk environments. But there are some causes of death for which there is good evidence of a direct causal impact. Cocaine carries a clear risk for cardiovascular events (e.g., heart attack, arrhythmia or stroke) (24). Risk of mortality can only be mitigated through reduced use. Yet, a major hurdle in reducing mortality associated with cocaine use is the lack of treatment options for cocaine dependence. There is no strong clinical evidence to support pharmacotherapies for cocaine dependence despite trials of various medicines(25-30). Psychosocial treatment (primarily contingency management) may decrease frequency of use and increase length of abstinence; longer-term impacts post-treatment are less clear (31). However, accessibility and affordability of psychosocial treatment are major issues in many countries, particularly for consumers of crack cocaine who are typically more socially marginalised relative to consumers of powder cocaine (32).

There are a range of secondary interventions that can address mortality risk pathways. Prevalence of HIV, HCV and other infectious diseases is often higher among people who use cocaine (and crack

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cocaine specifically) relative to the general population (33), driving AIDS-related mortality and likely contributing to excess mortality from liver disease. Regular HIV and HCV testing coupled with access to HIV antiretroviral therapies and HCV direct-acting antiviral agents could significantly reduce HIV and HCV related mortality. Access to these services is suboptimal in almost all countries (34). Strategies to prevent transmission include provision of sterile needles and smoking pipes, free condoms, and pre-exposure prophylaxis for HIV and sexually transmitted infections (25). Rigorous evaluation of some of these interventions is yet to be undertaken (particularly with respect to their impacts on mortality) however efforts to directly engage people who smoke crack cocaine in particular are laudable as existing services targeted at people who inject drugs may not meet the needs of this group.

Whilst preventable, reducing excess mortality due to traumatic causes (e.g., suicide, homicide, and accidental injury) is challenging. People with regular or problematic cocaine use are often from socioeconomically disadvantaged areas and are more likely to be exposed to environmental risk factors (e.g., violence, crime) than people who use other drugs (35). There is also a significant gap in evidence regarding interventions to reduce agitation related to stimulant intoxication and to manage violence risk more broadly amongst this group (36). Treatment for cocaine dependence thus needs to considerate of possible means for reducing the risk of injury and violent behaviour against a backdrop of broader environmental risk factors. CBT can also reduce suicide risk in substance-using populations (37).

Strengths and limitations

There has been a significant growth in the literature on mortality among people who report regular or problematic cocaine use, with 13 new cohorts identified since the previous review in 2008(8). Cohorts were recruited from regions with the highest levels of cocaine use (i.e., North America, Western Europe and Latin America; 1), with Australasia and Central Europe being notable exceptions. There are limited data on cocaine use in many of the remaining regions (e.g., Asian and African regions;

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1), highlighting the challenge of capturing health harms associated with regular or problematic cocaine use in these regions.

Limitations of existing studies included uncertain representativeness of samples and poor reporting of methodological detail. Many of the cohorts included here were derived from a single sample type: typically, those individuals engaged with health services. Linkage of administrative data from healthcare services could be used in future studies. Secondly, data were often missing on how causes of death were defined and, where available, varied between studies. Although there are likely many contributors to heterogeneity in cause-specific mortality estimates, use of standardised definitions of cause-specific mortality (38) would reduce this potential source of variability. Nonetheless, it is important to acknowledge that varying definitions of causes likely contributed to heterogeneity across studies. Ideally, an approach where people used standardised definitions of causes of mortality might occur in future studies (e.g., 38), permitting some examination of whether this explains some of the variation observed.

We were thorough in searching for relevant studies, however, we may have missed eligible cohorts or made errors. Where necessary, we sought additional data from authors and we generated estimates of CMRs, SMRs and RRs where possible. There are potential biases introduced by our computation of person years where this was not reported, including impaired capacity to account for censoring, however sensitivity analyses showed relative consistency between imputed and study-reported person-years when we tested our approach on those studies which did provide the latter information. Computing SMRs and RRs was achieved by using estimates from the GDB study, the most comprehensive effort globally to estimate prevalence of cocaine dependence and of population-level cause-specific mortality (14). Pooled SMRs including estimates imputed from GBD typically fell within the confidence interval of the pooled SMRs from author-reported estimates only. We acknowledge that the GBD data may have some limitations and there was minor variability between pooled

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estimates of author-reported SMRs versus author-reported and imputed SMRs. Reporting of all-cause and cause-specific SMRs in future work would improve confidence in existing estimates. Further, RRs should be treated with caution, being based on prevalence of cocaine dependence extracted from the GDB study, and likely of lower prevalence than regular and/or problematic cocaine use.

Finally, some variability in cohort definition and study design should also be noted. This was explored via meta-regression and stratified meta-analyses; although cohort definition (i.e. meeting criteria for dependence vs not) was not associated with heterogeneity in CMR/SMR, other aspects of study design (e.g., recruitment setting, sampling frame) were associated with heterogeneity in estimates.

Conclusions

There has been increased study of cohorts of people reporting regular or problematic cocaine use. For this reason, we could quantify excess mortality by cause of death in the current review. Synthesis of this evidence suggests people with regular or problematic cocaine use have, on average, six times (95%Cl four to nine times) higher mortality risk than their age and sex peers in the general population. Excess mortality risk is particularly evident for traumatic causes of deaths and causes likely attributable to infectious diseases. These deaths are mostly preventable. A lack of treatment options for cocaine dependence mean current efforts rely heavily on other prevention and intervention strategies to address risk pathways to mortality.

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Table 1: Characteristics of included studies

Study	Country	Years conducted ^a	Sample	N people	N person years ^b
1. Accurso, 2015(39)	United States	1990-2010	People "abusing" cocaine and presenting for detoxification at Chemical Dependence Unit at John Hopkins Bayview Medical Center in Baltimore between 1990-1991	315	5780*
2. Arendt, 2011(40)	Denmark	1996-2006	People receiving publicly funded treatment for illicit substance use disorder and primarily using cocaine across Denmark between 1996-2006, identified from the Danish Substance Abuse Treatment Register	838	2571*
3. Barrio, 2013(41)	Spain	2004-2010	People who reported regular cocaine use (≥ 52 days/last year) recruited from drug scenes and non-treatment settings in Madrid, Barcelona, and Seville, Spain, between 2004-2006	714	3922
4. Bohnert, 2017(23)	United States	2006-2011	People receiving Veteran Health Administration (VHA) care in the 2005 financial year diagnosed with cocaine use disorder and still alive in 2006 as identified using VHA National Patient Care Database	83808	468560*
5. Callaghan, 2013(22) (Callaghan, 2012(42))	United States	1990-2005	People hospitalised with a cocaine use disorder diagnosis in California between 1990-2005 from the Patient Discharge Database	48949	395738
6. de la Fuente, 2016(43) (Brugal, 2016(44); Colell, 2018(45); Molist, 2018(21))	Spain	1997-2008	People starting drug treatment for cocaine use disorder in a publicly owned or funded treatment centre in Barcelona and Madrid between 1997-2007	11905	65849*
7. Dias, 2011(46) ^c	Brazil	1992-2006	People who were consecutively admitted patients to Taipas General Hospital's inpatient treatment for crack/cocaine dependence between 1992- 1994	131	1182*
8. Gossop, 2002(47) ^c	United Kingdom	1995-1999	People who self-reported cocaine misuse recruited to treatment programs throughout England in 1995 as part of National Treatment Outcome Research Study (NTORS) cohort	227	926*
9. Hayashi, 2016(48) <i>(Tyndall, 2001(49))</i>	Canada	1996-2011	People who injected cocaine in the 6 months prior to enrolling in the open cohorts of Vancouver Injection Drug Users Study (VIDUS) and AIDS Care Cohort to Evaluate access to Survival Sciences (ACCESS) in Vancouver between 1996-2011	1719	11749
10. Hser, 2012(50)	United States	2002-2010	Mothers enrolled in a drug treatment program with cocaine as the primary drug of concern between 2000-2002 across California as identified through California Treatment Outcome Project (CalTOP)	511	5471*

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11. Lopez, 2004 (OFDT)(51) ^c	France	1992-2001	People arrested in 1992, 1993, 1996 and 1997 for cocaine/crack use/dealing as identified through database of police questioning files for use of narcotics	2212	11496
12. Markota, 2016(52)	United States	1999-2011	People aged 13-18 years old attending drug and alcohol treatment and with a positive cocaine urinary toxicology screen administered at clinical sites within Mayo Health Care System between 1999-2011	63	308*
13. Martell, 2009^(53)	United States	2003-2005	People with cocaine and opioid dependence enrolled in a randomised clinical trial for cocaine vaccine from greater New Haven between 2003-2005	114	58*
14. Nielsen, 2011(54)	Denmark	1999-2009	People diagnosed with cocaine abuse and at least one contact with a homeless shelter in Denmark between 1999-2009 as identified by the Danish Homeless Register	525	5362*
15. O'Driscoll, 2001(55)	United States	1994-1997	People who inject drugs within Seattle and King County and reported cocaine as their primary drug recruited between 1994-1996	340	931*
16. Pavarin, 2017(56) (Pavarin, 2008(57); Pavarin, 2013(58))	Italy	1989-2013	Individuals admitted to a public drug treatment for problems caused by primary use of cocaine in Bologna (North Italy) between 1989-2013	678	4753*
17. Ryb, 2009(59)	United States	1983-1997	People discharged from R Adams Cowley Shock Trauma Center with a positive cocaine urinary toxicology screen at admission between 1983-1995	2451	15932*
18. Sanvisens, 2014(60)	Spain	1985-2008	Patients admitted to hospital detoxification for primary cocaine abuse at one of three tertiary care facilities in Barcelona and the surrounding metropolitan area between 1985-2006	945	7155
19. van Haastrecht, 1996(61) ^c	Netherlands	1985-1993	People who are HIV+ and HIV- who self-reported injecting cocaine and were recruited in Amsterdam between 1985-1992 through "low-threshold" methadone programs and clinic workers for people who use drugs and engage in sex work.	632 ^d	194
20. Vlahov, 2008(62)	United States	1997-2002	People who inject drugs with self-reported injecting of cocaine everyday recruited from five U.S. cities between 1997-1999 through community-based outreach methods and enrolled in the second Collaborative Injection Drug Users Study (CIDUS-II)	102	486
21. Wang, 2005(63) ^c	United States	1988-2005	People who inject drugs who self-reported cocaine use in the previous 6 months, recruited through community in Baltimore between 1988-1989 and 1994-1998 and enrolled in AIDS Linked to Intravenous Experience study (ALIVE)	518#	3727

Note. *Italics* denotes associated secondary paper for the cohort; *Person years were not reported by study but calculated using formula within **Appendix H;** ^ Denotes that the study was a randomised controlled trial (RCT); # Denotes that the information was provided by the authors on request; ^a Period covers the start of recruitment until the end of follow-up; ^b Person-years were rounded to nearest whole number, though exact person-years reported were used for analysis; ^c Study was included in the previous review and information differs due to additional information being provided; ^d Study does not specify proportion of cohort that uses cocaine, but provides number of deaths within those who use cocaine.

Table 2: Summary of risk of bias and quality of study reporting

	Study risk o	f bias	Study reporting quality								
	Entire col	ort		Entire cohort ^a	Cocaine sub-cohort ^a						
	Representativeness	Outcome measurement	Cohort description (age/sex)	Completeness of data (numerator/denominator)	Cause of death definition ^b	Cohort description (age/sex)	Completeness of data (numerator/denominator)				
Accurso, 2015(39)	↑	\downarrow	↑	↓	NA	\downarrow	\downarrow				
Arendt, 2011(40)	\downarrow	\downarrow	1	1	NA	↓	\downarrow				
Barrio, 2013(41)	↓	↓ C	1	1	↓ ↓	1	↑				
Bohnert, 2017(23)	↑	\downarrow	1	\downarrow	1	↓	\downarrow				
Callaghan, 2013(22)	\downarrow	\downarrow	↑	↑	↑	↑	↑				
Callaghan, 2012(42)	↓	\downarrow	↑ C	• ↓	NA	1	\downarrow				
de la Fuente, 2016(43)	\downarrow	\downarrow	1	A	NA	1	\downarrow				
Brugal, 2016(44)	↓	¥	1	1	1	1	↑				
Colell, 2018(45)	↑	¥	1	↑ U	NA	1	↑				
Molist, 2018(21)	↓	¥	1	↑	1	1	↑				
Dias, 2011(46)	↑	1	1	\downarrow	+	1	Ļ				
Gossop, 2002(47)	↑	¥	\rightarrow	\downarrow	1	↓	Ļ				
Hayashi, 2016(48)	↓	¥	1	↑	1	¥	Ļ				
Tyndall, 2001(49)	↓	¥	1	↓	1	¥	Ļ				
Hser, 2012(50)	↓ ↓	\downarrow	↑	↓ ↓	↓	\downarrow	Ļ				
Lopez, 2004 (OFDT)(51)	↑	\downarrow	1	Ļ	\downarrow	↓	Ļ				
Markota, 2016(52)	↓ ↓	1	\downarrow	↓	Ļ	¥	Ļ				

				incluse and moreancy			
Martell, 2009^(53)	1	1	1	\downarrow	NA	1	\downarrow
Nielsen, 2011(54)	\downarrow	↓	1	\downarrow	1	↓	\downarrow
O'Driscoll, 2001(55)	\downarrow	Ļ	1	1	Ļ	↓	1
Pavarin, 2017(56)	↑	Ļ	1	\downarrow	¥	1	Ļ
Pavarin, 2008(57)	↑	Ļ	1	1	1	1	1
Pavarin, 2013(58)	1	Ļ	1	\downarrow	Ļ	1	↓
Ryb, 2009(59)	1	\downarrow	1	\downarrow	Ļ	1	Ļ
Sanvisens, 2014(60)	1	↓ ↓	1	1	1	1	1
van Haastrecht, 1996(61)	\downarrow	↓	↓)	1	↓	\downarrow	1
Vlahov, 2008(62)	↓	\downarrow		1	NA	\downarrow	\downarrow
Wang, 2005(63)	\downarrow	↓	1	↑	Ļ	↓	\downarrow
Higher risk of bias/ Lower quality of reporting ^c	9/21 (42.9%)	3/21 (14.9%)	3/21 (14.9%)	11/21 (52.4%)	10/17 (58.8%)	13/21 (61.9%)	16/21 (76.2%)

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Note. Full details of the risk of bias and quality of study reporting assessment are available in **Appendix H**, noting that studies with insufficient information to evaluate each criteria were deemed 'higher risk of bias'/'lower quality of reporting'; \downarrow denotes lower risk of bias or lower quality of reporting; \uparrow denotes higher risk of bias or higher quality of reporting^a 'Entire cohort' refers to reporting for sample of interest in the original study which may include people who do not consume cocaine; 'cocaine sub-cohort' refers to reporting where mortality outcomes were reported for the subsample of interest for this paper (i.e., people with regular/problematic cocaine use) but whom may not be the primary focus of the paper. ^b Includes 17 cohorts with cause-specific mortality data; 'NA' denotes cohorts where cause-specific mortality data were not collected; ^c Totals only include the primary publication if there are multiple publications reporting on the same cohort.

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Table 3: Pooled all-cause and cause-specific CMRs and SMRs for people with regular/problematic cocaine use, stratified by sex, age, drug use characteristics and region

	Crude m	ortality rat	e			Standardised mortality ratio				
	N cohort s	N people	Pooled crude mortality rate per 100PY (95%CI)	²	References	N cohorts (N author- reported SMR) ^e	N people	Pooled standardised mortality ratio (95%CI)	²	References
All-cause mortality									·	
Overall	16	69,954ª	1.24 (0.86-1.78)	98.8%	(22, 39-41, 43, 46, 48, 51, 52, 54-56, 60-63)	16 (6)	69,932ª	6.13 (4.15-9.05)	99.0%	(39-43, 46, 48, 51, 52, 54-56, 60-63)
Sex										
Women	6	25,217 ^b	0.66 (0.54-0.81)	55.6%	(22, 40, 43, 52, 54, 58)	6 (3)	25,202 ^{b,c}	4.59 (2.68-7.87)	94.2%	(40, 42, 43, 52, 54, 57)
Men	6	37,041 ^b	0.89 (0.50-1.56)	98.3%	(22, 40, 43, 52, 54, 58)	6 (4)	37,056 ^b	3.42 (2.86-4.10)	77.5%	(40, 42, 43, 52, 54, 57)
Age				Y						
< 30	2	3,677	0.67 (0.16-2.89)	97.1%	(21, 46)	2 (1)	3,677	7.75 (5.67-10.58)	0.0%	(21, 46)
≥ 30	3	8,358	1.01 (0.49-2.09)	81.5%	(21, 46, 57)	3 (1)	8,358	3.09 (1.77-5.37)	69.4%	(21, 46, 57)
Lifetime cocaine injection	5	2,679ª	3.60 (3.32-3.90)	88.4%	(48, 55, 61-63)	5 (0)	2,679ª	13.74 (12.67-14.90)	91.2%	(48, 55, 61-63)
Cocaine dependence/use disorder	8	64,286	1.09 (1.07-1.13)	98.4%	(22, 39, 40, 43, 46, 54, 56, 60)	8 (5)	64,286	3.24 (3.15-3.33)	97.6%	(22, 39, 40, 43, 46, 54, 56, 60)
GBD region*										
High-Income North America	7	52,006	1.56 (0.83-2.95)	99.2%	(22, 39, 48, 52, 55, 62, 63)	7 (1)	51,984	5.13 (2.34-11.25)	99.5%	(39, 42, 48, 52, 55, 62, 63)
Western Europe	8	17,817ª	0.93 (0.49-1.78)	98.6%	(40, 41, 43, 51, 54, 56, 60, 61)	8 (4)	17,817ª	6.01 (4.16-8.68)	94.6%	(40, 41, 43, 51, 54, 56, 60, 61)
Tropical Latin America	1	131	2.28 (1.57-3.33)	-	(46)	1 (1)	131	14.75 (9.92-21.17)	-	(46)
Cause-specific mortality										
Drug-related ^c	8	16,857	0.34 (0.10-1.15)	98.6%	(41, 44, 46, 47, 49, 55, 56, 63)	8 (2)	с	с	с	с
Suicide	8	100,854	0.07 (0.04-0.10)	72.6%	(21, 23, 41, 49, 52, 56, 59, 63)	8 (2)	100,854	6.26 (2.84-13.80)	94.4%	(21, 22, 46, 52, 56, 59)
Accidental injury	6	64,177	0.09 (0.04-0.22)	95.4%	(21, 22, 46, 52, 56, 59)	6 (3)	64,177	6.36 (4.18-9.68)	71.9%	(21, 42, 46, 52, 56, 59)
Cardiovascular disease	4	14,085	0.13 (0.07-0.24)	77.3%	(21, 49, 60, 63)	4 (1)	14,085	1.83 (0.39-8.57)	96.7%	(21, 49, 60, 63)
Homicide	3	14,487	0.09 (0.01-0.54)	98.4%	(21, 46, 59)	3 (1)	14,487	9.38 (3.45-25.48)	90.1%	(21, 46, 59)

AIDS-related	7	7,293	0.28 (0.12-0.63)	95.1%	(41, 45, 46, 49, 56, 60, 63)	6 (1)	6,576	23.12 (11.30-47.31) ^g	90.1%	(41, 45, 46, 56, 60, 63)
Cancer	5	14,763	0.11 (0.05-0.25)	87.1%	(21, 49, 56, 60, 63)	5 (2)	14,763	1.49 (0.70-3.16)	85.5%	(21, 49, 56, 60, 63)
Respiratory disease	5	6,217	0.09 (0.04-0.17)	63.2%	(41, 45, 49, 56, 63)	1 (1) ^f	852	24.12 (6.03-96.43)	-	(56)
Digestive diseases	3	13,140	0.14 (0.03-0.73)	96.1%	(21, 49, 63)	1 (1) ^f	11,905	1.90 (1.20-2.89)	-	(21)
Liver-related ^d	3	4,666	0.06 (0.01-0.55)	99.9%	(45, 46, 60)	3 (0)	4,666	3.36 (0.51-22.10)	92.5%	(45, 46, 60)

Cocaine use and mortality

Note.

* Regions are defined as per the Global Burden of Disease (GBD) project. No studies were found for the following GBD regions: Central Asia, Central Europe, Eastern Europe, Australasia, High-Income Asia Pacific, Southern Latin America, Andean Latin America, Caribbean, Central Latin America, North Africa & Middle East, South Asia, East Asia, Oceania, Southeast Asia, Central Sub-Saharan Africa, Eastern Sub-Saharan Africa, Southern Sub-Saharan Africa or Western Sub-Saharan Africa.

^a Except for van Haastrecht,1996(61), all studies reported the N of people who reported regular/problematic cocaine use.

^b Except for Arendt, 2011(40), all studies reported the N of people who reported regular/problematic cocaine use.

^c Though study estimates were available for SMR, these were deemed unstable and therefore not included. See **Appendix L** for the pooled estimate using author reported SMRs. It should be noted that drug poisoning deaths, for some studies, could include poisoning due to any drug and mental and behavioural disorders due to psychoactive substance use and other causes deemed by the authors to be drug-related (see Appendix J for ICD codes where reported).

^d There is partial overlap with digestive diseases deaths, but liver-related deaths includes those that were specified identified as liver related whereas digestive diseases deaths included any deaths within a broader definition encompassing the digestive system (i.e., Chapter 10 of ICD-10 codes).

^e The number in brackets denotes the number of cohorts where SMRs were reported by the authors, noting we imputed SMRs for those cohorts where these data were not reported (see Appendix L for pooled estimates using only author-reported SMRs.

^f Due to the nature of GBD classifications, expected number of deaths was unable to be estimated resulting in no imputed SMRs to be calculated.

^g Tyndall, 2001(49) excluded as estimate was not logical.

Table 4: Meta-regression of potential sources of heterogeneity in the pooled all-cause crude mortality rate (CMR) and standardised mortality ratio (SMR)

				Standardised mortality ratio				
	N studies	Coefficient (SE)	Adj. R²	Р	N studies	Coefficient (SE)	Adj. R²	Р
Sample characteristics at baseline	1					1	1	_
% Women	9	0.194 (0.543)	- 9.76%	0.576	9	0.038 (0.051)	39.96%	0.045
% Injecting	8	5.871 (3.272)	68.62%	0.019	8	3.091 (2.472)	17.35%	0.208
Type of cocaine use			- 1.54%				3.19%	
Cocaine/Cocaine and crack cocaine	16	ref			16	ref		
Speedball/Cocaine and Heroin	6	1.383 (0.546)		0.422	2	2.113 (1.400)		0.276
Geographic region			1.20%				- 7.77%	
% Western Europe	8	ref			8	ref		
% High-Income North America	7	1.689 (0.755)		0.262	7	- 0.207 (0.494)		0.682
% Tropical Latin America	1	2.447 (2.199)		0.337	1	0.844 (0.995)		0.412
Study characteristics					•			
Year of final follow-up	16	0.939 (0.034)	12.54%	0.103	16	0.925 (0.035)	16.88%	0.056
Sample size	15	1.000 (< 0.001)	- 7.82%	0.797	15	1.000 (< 0.001)	0.29%	0.345
Person years	16	1.014 (0.038)	- 6.45%	0.713	16	0.973 (0.038)	- 4.64%	0.502
Recruitment setting			-14.62%				20.51%	
Treatment clinics and other health	6	ref			6	ref		
services								
Hospital	3	0.750 (0.492)		0.664	3	0.247 (0.146)		0.034
Other	7	0.961 (0.489)		0.939	7	0.587 (0.268)		0.264
Sampling frame			34.92%				24.65%	
National/sub-national	9	ref			9	ref		
City	7	2.776 (0.985)		0.012	7	2.579 (1.042)		0.034

Note. Regions are defined as per the Global Burden of Disease (GBD) project. No studies were found for the following GBD regions: Central Asia, Central Europe, Eastern Europe, Australasia, High-Income Asia Pacific, Southern Latin America, Andean Latin America, Caribbean, Central Latin America, North Africa & Middle East, South Asia, East Asia, Oceania, Southeast Asia, Central Sub-Saharan Africa, Eastern Sub-Saharan Africa, Southern Sub-Saharan Africa or Western Sub-Saharan Africa. Ref: reference category; SE: standard error

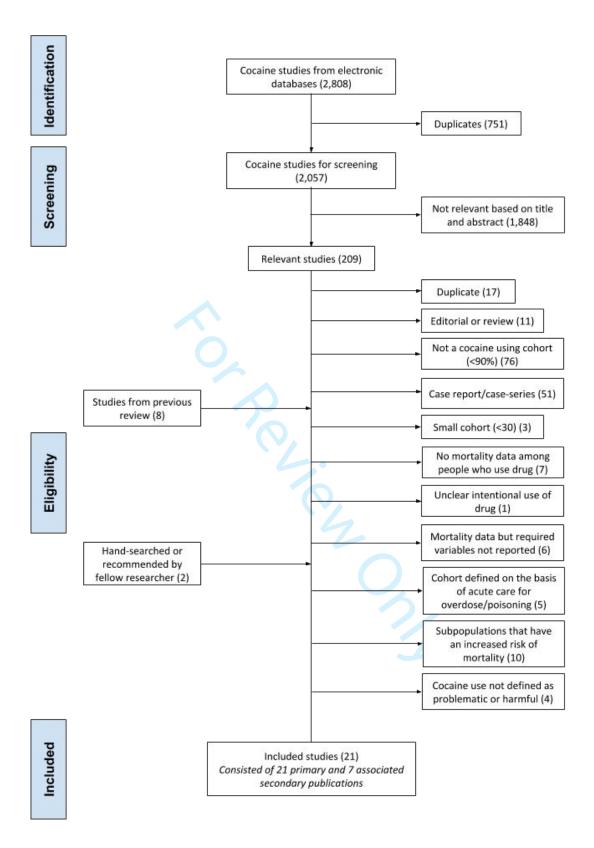


Figure 1. PRISMA flow diagram of studies reporting on mortality among people with regular or problematic cocaine use

Cocaine use and mortality

Figure 1. Pooled estimates of all-cause crude mortality rates (CMR) per 100 person-years (left) and standardised mortality ratio (SMR) (right) among people with regular

or problematic cocaine use, overall and by region

Study		%	Study		%
D	CMR/100PY (95% CI)	Weight	ID	SMR (95% CI)	Weigh
ligh-Income North America			High-Income North America		
/lahov, 2008 (United States)	0.62 (0.20, 1.92)	4.18	Accurso, 2015 (United States)*	2.11 (1.61, 2.77)	6.67
Accurso, 2015 (United States)	0.90 (0.69, 1.18)	6.71	Markota, 2016 (United States)*	2.32 (0.75, 7.18)	4.36
Aarkota, 2016 (United States)	0.98 (0.31, 3.02)	4.18	Callaghan, 2012 (United States)	2.96 (2.87, 3.05)	6.89
Callaghan, 2013 (United States)	1.10 (1.07, 1.13)	6.96	O'Driscoll, 2001 (United States)*	5.39 (3.19, 9.10)	6.13
D'Driscoll, 2001 (United States)	1.50 (0.89, 2.54)	6.10	Vlahov, 2008 (United States)*	5.48 (1.77, 17.00)	4.36
layashi, 2016 (Canada)	3.34 (3.02, 3.68)	6.93	Wang, 2005 (United States)*		6.82
Vang, 2005 (United States)	➡ 4.69 (4.03, 5.45)	6.89	Hayashi, 2016 (Canada)*	♦ 14.05 (12.72, 15.51)	6.86
Subtotal (I-squared = 99.2%, p = 0.000)	1.56 (0.83, 2.95)	41.95	Subtotal (I-squared = 99.5%, p = 0.000)	5.13 (2.34, 11.25)	42.10
Vestern Europe			Western Europe		
Barrio, 2013 (Spain)	0.23 (0.12, 0.44)	5.71	Lopez, 2004 (France)*	2.09 (1.68, 2.61)	6.75
Pavarin, 2017 (Italy)	0.53 (0.36, 0.78)	6.45	Nielsen, 2011 (Denmark)* →	3.98 (3.17, 4.99)	6.74
le la Fuente, 2014 (Spain)	0.53 (0.48, 0.59)	6.93	Barrio, 2013 (Spain)	4.70 (2.40, 9.00)	5.75
opez, 2004 (France)	0.70 (0.56, 0.87)	6.80	de la Fuente, 2014 (Spain)	4.90 (4.40, 5.40)	6.86
Arendt, 2011 (Denmark)	0.70 (0.44, 1.11)	6.27	Pavarin, 2017 (Italy)	5.21 (3.52, 7.71)	6.44
Vielsen, 2011 (Denmark)	1.40 (1.12, 1.75)	6.79	Sanvisens, 2014 (Spain)*	6.12 (5.47, 7.00)	6.85
Sanvisens, 2014 (Spain)	2.80 (2.50, 3.20)	6.91	Arendt, 2011 (Denmark)	6.40 (3.90, 10.00)	6.26
an Haastrecht, 1996 (Netherlands)	4.64 (2.41, 8.92)	5.70	van Haastrecht, 1996 (Netherlands)*	84.93 (44.19, 163.24)	5.77
Subtotal (I-squared = 98.6%, p = 0.000)	0.93 (0.49, 1.78)	51.56	Subtotal (I-squared = 95.3%, p = 0.000)	6.01 (4.16, 8.68)	51.43
Tropical Latin America			Tropical Latin America		
Dias, 2011 (Brazil)	2.29 (1.57, 3.33)	6.49	Dias, 2011 (Brazil)	14.75 (9.92, 21.17)	6.47
Subtotal (I-squared = .%, p = .)	2.28 (1.57, 3.33)	6.49	Subtotal (I-squared = .%, p = .)	14.75 (10.10, 21.55)	6.47
Overall (I-squared = 98.8%, p = 0.000)	1.24 (0.86, 1.78)	100.00	Overall (I-squared = 99.0%, p = 0.000)	6.13 (4.15, 9.05)	100.0
· · · · · · _		<u></u> 2	i		

Cocaine use and mortality

Figure 3: Pooled estimates of cause-specific crude mortality rates (CMR) per 100 person-years (left) and standardised mortality ratio (SMR) (right) among people with

regular or problematic cocaine use

idy	CMR/100PY (95% CI)	% Weight	Study ID	CMR/100PY (95% CI)	% Weight
ug-Related varin, 2017 (Italy)	0.06 (0.02, 0.20)	11.77	AIDS-Related		
rrio, 2013 (Spain)	0.08 (0.02, 0.20)	11.77	Colell, 2018 (Spain)	0.02 (0.01, 0.04)	13.91
ugal, 2016 (Spain)	0.08 (0.06, 0.09)	13.08	Pavarin, 2017 (Italy)	0.02 (0.00, 0.15)	8.75
as, 2011 (Brazil)	0.25 (0.08, 0.79)	11.77	Barrio, 2013 (Spain)	0.05 (0.01, 0.20)	11.62
ang, 2005 (United States)		12.96	Dias, 2011 (Brazil)	0.51 (0.23, 1.13)	14.88
ssop, 2002 (United Kingdom)	0.97 (0.51, 1.87)	12.64	Sanvisens, 2014 (Spain)	 0.82 (0.64, 1.06) 	17.03
Driscoll, 2001 (United States)	0.99 (0.69, 1.41)	12.97	Tyndall, 2001 (Canada)	1.15 (0.80, 1.64)	16.77
ndall, 2001 (Canada)	→ 2.07 (1.58, 2.70)	13.03	Wang, 2005 (United States)	➡ 1.66 (1.30, 2.13)	17.04
btotal (I-squared = 98.6%, p = 0.000)	0.34 (0.10, 1.15)	100.00	Subtotal (I-squared = 95.1%, p = 0.000)	0.28 (0.12, 0.63)	100.00
icide rrio, 2013 (Spain)	0.03 (0.00, 0.18)	4.15	Cancer		
b, 2009 (United States)	0.03 (0.01, 0.08)	12.49	Molist, 2018 (Spain)	0.06 (0.04, 0.08)	23.79
hnert, 2017 (United States)	0.05 (0.04, 0.06)	24.54	Pavarin, 2017 (Italy)	0.06 (0.02, 0.20)	16.93
list, 2018 (Spain)	0.05 (0.04, 0.07)	21.84	Sanvisens, 2014 (Spain)	0.08 (0.04, 0.19)	20.06
ing, 2005 (United States)	0.05 (0.01, 0.21)	7.13	Tyndall, 2001 (Canada)	0.11 (0.04, 0.36)	16.93
varin, 2017 (Italy)	0.08 (0.03, 0.22)	11.10	Wang, 2005 (United States)	0.35 (0.20, 0.60)	22.28
ndall, 2001 (Canada)	0.27 (0.13, 0.56)	14.58	Subtotal (I-squared = 87.1%, p = 0.000)	0.11 (0.05, 0.25)	100.00
Inkota, 2016 (United States)	• 0.33 (0.05, 2.31)	4.15			
btotal (I-squared = 72.6%, p = 0.001)	0.07 (0.04, 0.10)	100.00	Respiratory Disease	the second second second second second second	
cidental Injury			Pavarin, 2017 (Italy)	0.04 (0.01, 0.17)	13.82
llaghan, 2013 (United States)	0.03 (0.02, 0.03)	20.98	Barrio, 2013 (Spain)	0.05 (0.01, 0.20)	13.82
list, 2018 (Spain)	0.07 (0.05, 0.09)	20.69	Colell, 2018 (Spain)	0.06 (0.04, 0.10)	29.84
as, 2011 (Brazil)	0.08 (0.01, 0.60)	9.99	Tyndall, 2001 (Canada)	0.11 (0.04, 0.36)	17.30
varin, 2017 (Italy)	0.13 (0.06, 0.28)	17.86 20.49	Wang, 2005 (United States)	0.21 (0.11, 0.43)	25.22
rkota, 2016 (United States)	0.33 (0.05, 2.31)	9.99	Subtotal (I-squared = 63.2%, p = 0.028)	0.09 (0.04, 0.17)	100.00
btotal (I-squared = 95.4%, p = 0.000)	0.09 (0.04, 0.22)	100.00			
	0.00 (0.01, 0.22)	100.00	Digestive Disease		
rdiovascular Disease			Molist, 2018 (Spain)	0.03 (0.02, 0.05)	33.92
list, 2018 (Spain)	0.07 (0.05, 0.09)	31.27	Wang, 2005 (United States)	0.27 (0.14, 0.50)	33.13
nvisens, 2014 (Spain)	0.13 (0.07, 0.24)	24.47	Tyndall, 2001 (Canada)	0.34 (0.18, 0.66)	32.96
ing, 2005 (United States)	0.19 (0.09, 0.39)	22.71	Subtotal (I-squared = 96.1%, p = 0.000)	0.14 (0.03, 0.73)	100.00
ndall, 2001 (Canada) = 0.004)	0.23 (0.10, 0.51)	21.55			
Diotal (1-Squared - 11.3%, p - 0.004)	0.13 (0.07, 0.24)	100.00	Liver-Related	11 March 10 Aug March 201	
micide			Colell, 2018 (Spain)	0.01 (0.01, 0.03)	35.20
list, 2018 (Spain)	0.01 (0.00, 0.02)	32.87	Dias, 2011 (Brazil)	0.08 (0.01, 0.60)	28.55
b, 2009 (United States)	0.08 (0.04, 0.13)	33.51	Sanvisens, 2014 (Spain)	0.22 (0.14, 0.37)	36.25
as, 2011 (Brazil)	1.35 (0.83, 2.21)	33.62	Subtotal (I-squared = 94.6%, p = 0.000)	0.06 (0.01, 0.55)	100.00
btotal (I-squared = 98.4%, p = 0.000)	0.09 (0.01, 1.54)	100.00	Andre and and a second and a		
.001 .01	5 1 2 4		.001 .01 .5	1 2 4	

Cocaine use and mortality

Figure 4: Pooled estimates of cause-specific standardised mortality ratio (SMR) (right) among people with regular or problematic cocaine use

uicide yb, 2009 (United States)* ohnert, 2017 (United States)* /ang, 2005 (United States)* avarin, 2013 (Spain)* avarin, 2017 (Italy) tolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	SMR (95% CI) 1.86 (0.77, 4.46) 3.06 (2.69, 3.48) 3.15 (0.79, 12.58) 4.19 (0.59, 29.75) 8.33 (3.13, 22.19) 14.86 (11.21, 19.33)	Weight 13.54 16.17 10.84	ID AIDS-Related Colell, 2018 (Spain)* Pavarin, 2017 (Italy)	SMR (95% CI) 2.85 (1.07, 7.58)	Weigh
yb, 2009 (United States)* ohnert, 2017 (United States)* /ang, 2005 (United States)* avarin, 2013 (Spain)* avarin, 2017 (Italy) lolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	3.06 (2.69, 3.48) 3.15 (0.79, 12.58) 4.19 (0.59, 29.75) 8.33 (3.13, 22.19)	16.17 10.84	Colell, 2018 (Spain)*	2.85 (1.07, 7.58)	16 15
ohnert, 2017 (United States)* /ang, 2005 (United States)* arrio, 2013 (Spain)* avarin, 2017 (Italy) lolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	3.06 (2.69, 3.48) 3.15 (0.79, 12.58) 4.19 (0.59, 29.75) 8.33 (3.13, 22.19)	16.17 10.84		2.85 (1.07, 7.58)	16 15
/ang, 2005 (United States)* arrio, 2013 (Spain)* avarin, 2017 (Italy) lolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	3.15 (0.79, 12.58) 4.19 (0.59, 29.75) 8.33 (3.13, 22.19)	10.84	Pavarin 2017 (Italy)		10.15
arrio, 2013 (Spain)* avarin, 2017 (Italy) lolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	4.19 (0.59, 29.75) 8.33 (3.13, 22.19)			4.98 (0.70, 35.34)	8.46
avarin, 2017 (Italy) lolist, 2018 (Spain) yndall, 2001 (Canada)* larkota, 2016 (United States)*	8.33 (3.13, 22.19)		Sanvisens, 2014 (Spain)*	➔ 30.28 (23.46, 39.08)	22.49
Iolist, 2018 (Spain) yndall, 2001 (Canada)* Iarkota, 2016 (United States)*		8.14	Dias, 2011 (Brazil)*	38.12 (17.13, 84.86)	17.96
yndall, 2001 (Canada)*	14 86 (11 21 19 33)	13.01	Barrio, 2013 (Spain)*	47.28 (11.83, 189.07)	
larkota, 2016 (United States)*		15.94	Wang, 2005 (United States)*	← 63.87 (49.79, 81.92)	22.53
	15.12 (7.21, 31.71)	14.22	Subtotal (I-squared = 90.1%, p = 0.000)	23.12 (11.30, 47.31)	100.00
	18.20 (2.56, 129.22)		Subiotal (1-squared = 50.1%, p = 0.000)	25.12 (11.50, 47.51)	100.0
ubtotal (I-squared = 94.4%, p = 0.000)	6.26 (2.84, 13.80)	100.00	Ganaar		
Description and the second s Second second seco			Cancer	0.54 (0.04, 4.40)	40.00
ccidental Injury			Sanvisens, 2014 (Spain)*	0.54 (0.24, 1.19)	19.89
ias, 2011 (Brazil)*	1.43 (0.20, 10.17)	4.00	Molist, 2018 (Spain)	0.95 (0.74, 1.19)	25.01
allaghan, 2013 (United States)	3.80 (2.30, 5.30)	23.62	Pavarin, 2017 (Italy)	1.98 (0.64, 6.15)	16.24
yb, 2009 (United States)*	5.72 (4.04, 8.08)	25.47	Tyndall, 2001 (Canada)*	2.02 (0.65, 6.25)	16.24
Iolist, 2018 (Spain)	8.51 (6.74, 10.50)	28.36	Wang, 2005 (United States)*	3.93 (2.28, 6.76)	22.62
avarin, 2017 (Italy)	9.93 (4.46, 22.09)	14.55	Subtotal (I-squared = 85.5%, p = 0.000)	1.49 (0.70, 3.16)	100.0
larkota, 2016 (United States)*	29.49 (4.15, 209.33)	4.00	1 KOREL MEDINE INVESTIGATION OF THE REAL PROPERTY O		
ubtotal (I-squared = 71.9%, p = 0.003)	6.36 (4.18, 9.68)	100.00	Respiratory Disease		
			Pavarin, 2017 (Italy)	◆ 24.12 (6.03, 96.43)	100.0
ardiovascular Disease			Subtotal (I-squared = .%, p = .)	24.12 (6.03, 96.45)	100.0
yndall, 2001 (Canada)*	0.11 (0.05, 0.24)	24.41		2	100.0
/ang, 2005 (United States)*	1.95 (0.93, 4.09)	24.64	Digestive Disease		
Iolist, 2018 (Spain)	2.97 (2.35, 3.67)	25.98	Molist, 2018 (Spain)	1.90 (1.20, 2.89)	100.00
anvisens, 2014 (Spain)*	16.29 (8.48, 31.31)	24.96	Subtotal (I-squared = .%, p = .)	1.90 (1.20, 2.89)	100.0
ubtotal (I-squared = 96.7%, p = 0.000)	1.83 (0.39, 8.57)	100.00	Subiolai (I-squaleu – .7%, p – .)	1.90 (1.23, 2.96)	100.0
			in a state		
omicide			Liver-Related		
yb, 2009 (United States)*	5.13 (2.91, 9.04)	33.08	Colell, 2018 (Spain)*	0.71 (0.32, 1.58)	35.74
Iolist, 2018 (Spain)	6.42 (3.43, 10.88)	32.94	Dias, 2011 (Brazil)*	7.58 (1.07, 53.85)	27.02
ias, 2011 (Brazil)*	24.39 (14.94, 39.81)	33.98	Sanvisens, 2014 (Spain)*	8.31 (5.09, 13.56)	37.24
ubtotal (I-squared = 90.1%, p = 0.000)	9.38 (3.45, 25.48)	100.00	Subtotal (I-squared = 92.5%, p = 0.000)	3.36 (0.51, 22.10)	100.0

* Imputed SMR was calculated for the study

Webappendices

"All-cause and cause-specific mortality among people with regular or problematic cocaine use: A systematic review and meta-analysis"

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Appendix A: PRISMA checklist

Section/topic	#	Checklist item	Reported on page #
TITLE			
Title	1	Identify the report as a systematic review, meta-analysis, or both.	1
ABSTRACT			
Structured summary	2	Provide a structured summary including, as applicable: background; objectives; data sources; study eligibility criteria, participants, and interventions; study appraisal and synthesis methods; results; limitations; conclusions and implications of key findings; systematic review registration number.	3-4
INTRODUCTION			
Rationale	3	Describe the rationale for the review in the context of what is already known.	5
Objectives	4	Provide an explicit statement of questions being addressed with reference to participants, interventions, comparisons, outcomes, and study design (PICOS).	6
METHODS	•		
Protocol and registration	5	Indicate if a review protocol exists, if and where it can be accessed (e.g., Web address), and, if available, provide registration information including registration number.	7; Appendix B p.6-7
Eligibility criteria	6	Specify study characteristics (e.g., PICOS, length of follow-up) and report characteristics (e.g., years considered, language, publication status) used as criteria for eligibility, giving rationale.	7-8; Appendix E pp. 10-11
Information sources	7	Describe all information sources (e.g., databases with dates of coverage, contact with study authors to identify additional studies) in the search and date last searched.	7; Appendix C p. 9
Search	8	Present full electronic search strategy for at least one database, including any limits used, such that it could be repeated.	Appendix D p. 9
Study selection	9	State the process for selecting studies (i.e., screening, eligibility, included in systematic review, and, if applicable, included in the meta-analysis).	7-8
Data collection process	10	Describe method of data extraction from reports (e.g., piloted forms, independently, in duplicate) and any processes for obtaining and confirming data from investigators.	8
Data items	11	List and define all variables for which data were sought (e.g., PICOS, funding sources) and any assumptions and simplifications made.	9-10
Risk of bias in individual studies	12	Describe methods used for assessing risk of bias of individual studies (including specification of whether this was done at the study or outcome level), and how this information is to be used in any data synthesis.	9; Appendix G pp. 15-16
Summary measures	13	State the principal summary measures (e.g., risk ratio, difference in means).	9-10

Synthesis of results	14	Describe the methods of handling data and combining results of studies, if done, including measures of consistency (e.g., I ²) for each meta-analysis.	10
Section/topic	#	Checklist item	Reported on page #
Risk of bias across studies	15	Specify any assessment of risk of bias that may affect the cumulative evidence (e.g., publication bias, selective reporting within studies).	9; Appendix J
Additional analyses	16	Describe methods of additional analyses (e.g., sensitivity or subgroup analyses, meta-regression), if done, indicating which were pre-specified.	10
RESULTS			
Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the review, with reasons for exclusions at each stage, ideally with a flow diagram.	Figure 1
Study characteristics	18	For each study, present characteristics for which data were extracted (e.g., study size, PICOS, follow-up period) and provide the citations.	Table 1; Appendix J
Risk of bias within studies	19	Present data on risk of bias of each study and, if available, any outcome level assessment (see item 12).	12-16; Table 2
Results of individual studies	20	For all outcomes considered (benefits or harms), present, for each study: (a) simple summary data for each intervention group (b) effect estimates and confidence intervals, ideally with a forest plot.	Figure 2-3; Appendix J
Synthesis of results	21	Present results of each meta-analysis done, including confidence intervals and measures of consistency.	16-23; Table 3; Figure 2-3
Risk of bias across studies	22	Present results of any assessment of risk of bias across studies (see Item 15).	Table 2
Additional analysis	23	Give results of additional analyses, if done (e.g., sensitivity or subgroup analyses, meta-regression [see Item 16]).	Table 4
DISCUSSION	I		
Summary of evidence	24	Summarize the main findings including the strength of evidence for each main outcome; consider their relevance to key groups (e.g., healthcare providers, users, and policy makers).	24-25
Limitations	25	Discuss limitations at study and outcome level (e.g., risk of bias), and at review-level (e.g., incomplete retrieval of identified research, reporting bias).	25-26
Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.	27
FUNDING			
Funding	27	Describe sources of funding for the systematic review and other support (e.g., supply of data); role of funders for the systematic review.	28

Appendix B: Description of methods in PROSPERO (ID: CRD42018094623)

The aim of this project will be an update of global reviews of available data on all-cause and causespecific mortality among people who use cocaine. The previous reviews include studies published from 1980 until 2008 new searches will be undertaken to include studies published from 2009 to current.

The Medline, Embase and PsycINFO peer-reviewed literature databases will be searched using the OVID[™] interface/platform for relevant articles published from the time-periods 2009 till current. Articles of interest comprise those likely to contain data describing all-cause and cause-specific crude mortality rates (CMR) and/or standardised mortality ratios (SMR) among people who use cocaine. Sets of search strings incorporating both keywords and Medical Subject Headings (MeSH terms) reflecting drug type and mortality epidemiology from the previous reviews will be revised and expanded for this updated search. Searches will be limited to human literature. No other restrictions will be applied to the search; citations for papers in languages other than English will be included and read via Google Translate. Citations from these searches will be imported into an Endnote[™] library, and duplicate citations removed.

To check for missing peer-review literature, reference lists for relevant systematic reviews identified in the peer-review literature search will be hand searched for additional papers not already identified. A final list of included studies will be distributed to experts to check if any relevant studies are missed.

Each set of search results (title and abstract) will first be screened by one team member in Covidence. All papers marked as excluded will be reviewed by a second person to ensure accuracy in first-pass screening. Each study for full-text screening will be reviewed in full by two people. Conflicts will be resolved through discussion and referral to a third party if needed.

Data will be independently extracted into an Excel worksheet template by one member of the research team and checked by a second member of the team. Bibliographic information will be extracted in addition to study specific information. Data entry will be standardized by use of a manual, which contains data entry rules. Where data are incomplete, authors will be contacted via email to obtain additional information.

Variables extracted will include study information and sample information (treatment status, HIV status, sex, percentage of sample injecting, treatment engagement). Crude mortality rates (CMRs) and standardized mortality ratios (SMRs) will be extracted as mortality measures. Cause of death information will be extracted for AIDS, overdose, suicide, traumatic (accident, homicide, injury, violence and poisoning), disease-related deaths, and other key categories. Disease-related deaths will be recoded according to the following categories; cardiovascular (endocarditis, myocardial infarction, circulatory system disease and cardiovascular disease), cerebrovascular, respiratory (pneumonia and chronic respiratory disease), liver (cirrhosis, viral hepatitis and liver disease), cancer (neoplasm, tumour and carcinoma), digestive (digestive system disease, nephritis and haemorrhage from duodenal ulcer), nervous system disease and other diseases (tuberculosis, bacterial infection, unspecified 'other disease' or 'natural causes' and when disease categories are not separated).

The quality index used in the previous reviews will be adopted for consistency. The quality index assesses each study on nine individual criteria: case ascertainment, measurement, diagnosis, estimate, numerator and denominator, data catchment, completeness, representativeness and age/sex variables. Each criterion will include a rating scale and the individual scores tallied to provide

an overall quality score. Study information necessary for quality assessment will be extracted to the Excel template. The greater the quality score, the higher the methodological quality of the study. Mortality estimates from higher rating studies may be given additional weighting in the calculation of final estimates. Quality estimates may also be used to undertake sensitivity analyses. Crude mortality rates will be calculated as per 100 person-years. Where person-years are not reported nor made available by the authors, an approximate person-year of follow-up will be calculated, with the assumption that deaths occurred halfway through the follow-up period, so that each case contributes half the person-year follow-up of survivors.

Estimates will be analysed in subgroups according to population type (for example, estimates from cohorts recruited on the basis of having a chronic physical health condition associated with a high mortality rate may be considered separately). Random-effects meta-analyses to determine pooled all-cause and cause-specific CMR and SMR estimates will be performed using STATA. This approach uses inverse variance weighting to calculate: fixed- and random effects pooled summary estimates; confidence limits; a test for differences between study effects; and an estimate of between-study variance. The random effects model allows for heterogeneity between as well as within studies; expecting high levels of heterogeneity between cohorts, the random-effects model will be used in all meta-analyses, with confirmation through the heterogeneity c2 and I-squared statistic. To investigate the source of this heterogeneity in an attempt to reduce it, cohorts may be divided into subsamples (e.g., by sex, age group, treatment status and HIV status) and/or these factors studied as possible risk factors for mortality via meta-regression in Stata.

4.ex Only

Medline	Compiled by the U.S. National Library of Medicine (NLM) and published on the Web by Community of Science, MEDLINE [®] is the world's most comprehensive source of life sciences and biomedical bibliographic information. It contains nearly eleven
	million records from over 7,300 different <u>publications</u> from 1965 to November 16, 2005.
	(Source: <u>http://medline.cos.com/docs/abmedl.shtml</u>)
EMBASE	EMBASE is a biomedical and pharmacological database.
	The EMBASE journal collection is international with over 5,000 biomedical journals from 70 countries. EMBASE contains over 11 million records from 1974 to present. EMBASE features comprehensive coverage of:
	 Drug Research, Pharmacology, Pharmacy, Pharmacoeconomics, Pharmaceutics and Toxicology
	 Human Medicine (Clinical and Experimental)
	Basic Biological Research
	Health Policy and Management
	Public, Occupational and Environmental Health
	Substance Dependence and Abuse
	• Psychiatry
	Forensic Science Diamodical Engineering and Instrumentation
	 Biomedical Engineering and Instrumentation
	(Source:
	http://www.elsevier.com/wps/find/bibliographicdatabasedescription.cws_home/523
	328/description#description
PsychINFO	PsycINFO is an abstract database of psychological literature from the 1800s to the present. More than 2.4 million records as of January 2008, including journals, books and dissertations. Over 2150 journal titles covered, 98% peer-reviewed; aso books and dissertations.
	(Source: <u>http://www.apa.org/psycinfo/</u>)

Appendix C: Description of databases used

Information

Database

Appendix D: Search strings and number of results for electronic literature searches

Database	Search group	Search terms			
Medline*	Cocaine	Cocaine			
		exp Cocaine/ or exp Co Disorders/ or exp Coca		-	Related
	Mortality	Mortal\$ or fatal\$ or de	eath\$		
		exp DEATH/ or exp "C exp Fatal Outcome/ or		-	TH, SUDDEN/ or
EMBASE#	Cocaine	Cocaine			
		exp Cocaine/ or exp Co Derivative/	ocaine Depenc	lence/ or exp Co	ocaine
	Mortality	Mortal* or fatal* or de	eath*		
		exp DEATH/ or exp "C or exp SUDDEN DEATH		· ·	
PsychINFO^	Cocaine	Cocaine			
		exp Cocaine/ or exp C	rack Cocaine/		
	Mortality	Mortal* or fatal* or de	eath*		
		exp "DEATH AND DYIN	NG"/ or exp Mo	ortality/ or exp	Mortality Rate
* 'key-words' in lo	owercase, 'MeSH' terms	in bold # 'key-words' in	lowercase, 'EMTREE	e' terms in bold	
^ 'key words' in lo	wercase, explode terms	in bold			
	Search to	erms		Database	
		-	EMBASE	Medline	PsycINFO

Appendix E: Inclusion and exclusion criteria for studies

Inclusion criteria for the study population comprised:

- Cohort studies of people using cocaine (where ≥90% of the sample report use of the substance of interest; people may report using multiple drugs);
- Cohort studies that are based on a sub-group of people who use cocaine (e.g. people who use cocaine who are incarcerated).
- Cohort studies that are not based on drug use (e.g. a cohort of homeless people), but a cocaine-using sub-group is identified, and results are presented separately for this sub-group.
- Case-control studies where cases are defined by cocaine use and mortality is reported for cases and controls separately
- Clinical trials of interventions to treat cocaine use disorder (randomised controlled trials; nonrandomised trials) where mortality may be reported at follow-up. Interventions may include pharmacotherapies or non-pharmacological treatments
- Secondary publications of an included cohort where the outcomes are not mortality (these may be used to assess risk of bias or to extract other baseline/demographic data not reported in the primary publication).

Inclusion criteria for the study outcome comprised of:

• Reporting raw mortality data (all-cause or cause-specific) or studies where such data can be obtained via contacting study authors.

Exclusion criteria comprised of:

- Cohort studies with a sample of less than 30 participants.
- Case reports/case series of drug-related deaths.
- Works that do not present original data (e.g. letters to the editor, editorials, commentaries)
- Works that do not present mortality data and where data could not be obtained via contacting study authors.
- Systematic reviews (although identified relevant reviews were handsearched for relevant studies which will be assessed independently against the inclusion/exclusion criteria)
- Studies where use of cocaine and mortality are not reported for the same sample of people
- Studies where it is unclear that everyone in the cohort reported intentional use of cocaine.
- Clinical trials of interventions for non-substance use disorders that are within a cocaine cohort e.g. clinical trials of HCV medicines in a cocaine cohort.

- Studies where the cohort was defined on the basis of acute care for overdose/poisoning.
- Studies where the cohort was defined by a population that has increased risk of mortality (e.g., severe health presentations or conditions, cardiovascular presentations related to cocaine exposure and being HIV positive)
- Studies where cohort reports cocaine use but it is not defined as regular or problematic.

FOR REVIEW ONLY

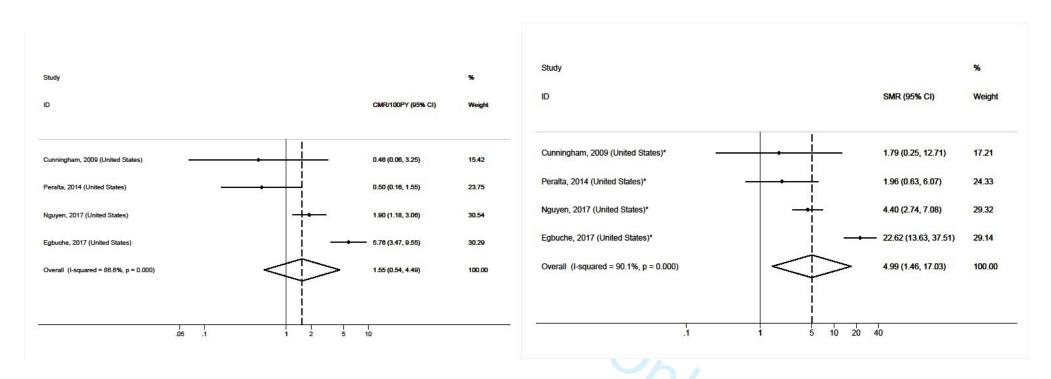
Appendix F: Cohort studies not included in main analyses

Table F1. List of identified cohort subpopulations not included in the main analyses (n=7)

Study		Country	Years conducted	Sample	N	Person years follow up	% women	Age at baseline	% PWID
1.	Cunningham, 2009 ¹	United States	2002-2005	People who presented for cocaine-induced chest pains, identified through a positive cocaine urine screen, to the Hurley Medical Center between 2002-2004 and followed for 1 year	219	218.5*	35%		
2.	Egbuche, 2017 ²	United States	2011-2014	People with heart failure presenting to Grady Memorial Hospital, Atlanta with a positive cocaine urine toxicology test between 2011-2014	268	260.5*	24.25%	54 (mean)	
3.	Nguyen, 2017 ³	United States	1993-2012	People hospitalised for acute decompensated heart failure at Michael E. DeBakey Veterans Affairs Medical Center, Houston between 1993-2012 with ICD coding of cocaine use and positive cocaine urine toxicology screen	90	892.539*	1.1%	55.5	
4.	Peralta, 2014 ^{^4}	United States	NR	People with cocaine-induced chest pain randomised to either coronary computed tomographic angiography or standard of care protocol	202	601.5*			
5.	Campa, 2016⁵	United States	NR	People with HIV from the Miami Adult Studies on HIV (MASH) cohort who self-reported cocaine use and/or had a positive cocaine urine toxicology screen	156	670.8*	31%	46.4	
6.	Nacher, 2009 ⁶	French Guiana	1996-2007	People with HIV admitted to Cayenne, Kourou, and Saint Laurent du Maroni Hospitals with crack cocaine dependence between 1996-2007 in French Hospital Database for HIV	76				
7.	Weber, 2015 ⁷	Switzerland	2007-2013	People with HIV registered prior to April 2007 treated at outpatient clinics at five university hospitals and two large district hospitals and who self-reported cocaine use	947#	4531.4*			32.73%

Note. *PYFU were not reported by study but calculated using formula within **Appendix E;** ^ denotes that the study was a randomised controlled trial (RCT); NR denotes that variable was not reported by the study; # denotes that the information was additionally provided by authors

Figure F1. Forest plots displaying the pooled all-cause CMR per 100 person-years (left) and all-cause SMR (right) with studies containing cohorts based on cardiovascular presentations.



Study % ID CMR/100PY (95% CI) Weigh Weber, 2015 (Switzerland) 1.70 (1.16, 2.48) 65.80 Colell, 2018 (Spain) 1.87 (0.97, 3.27) 25.67 Campa, 2016 (United States) 1.94 (0.68, 5.56) 8.54 Overall (I-squared = 0.0%, p = 0.950) 1.76 (1.29, 2.40) 100.00 4 .5 2 6 1

Figure F2. Forest plot displaying the pooled all-cause CMR per 100 person-years with studies that reported HIV positive subgroups.

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Appendix G: Quality of reporting and risk of bias assessment tool

	Description	Options	Rating
Representativeness ^a	How representative of the population are the cohort?	One sample type, one location	Higher risk of bias
		One sample type, multiple locations	Reviewer judgement required - Lower risk of bias OR higher risk of bias
		Multiple sample types, one location	Lower risk of bias
		Multiple sample types, multiple locations	Lower risk of bias
		Insufficient information to determine	
			Higher risk of bias
Outcome measurement	How was mortality measured?	Death registry or death certificate data	Lower risk of bias
		Medical record, personal contacts, or other source	Higher risk of bias
		Insufficient information to determine	Higher risk of bias
QUALITY OF REPORTIN	G DOMAINS		
	Description	Options	Rating
Completeness of	Was the cohort adequately described in	No – study did not report age or sex data for sample	Lower quality of reporting
cohort	terms of age (mean/median, or		
characterisation	categorical age data) and sex?		
		Yes – Age and sex data for the sample was reported	Higher quality of reporting
Completeness of outcome data	Did the study report both numerator and denominator for main outcome analysis?	No – Numerator, denominator, or both missing	Lower quality of reporting
		Yes – numerator and denominator reported for main outcome analysis	Higher quality of reporting
Definitions use for cause-specific deaths reported	Did the study report the definitions used for cause-specific deaths i.e. ICD codes	N/A – the study did not report cause-specific deaths	N/A
		No – no definitions provided (e.g. the study states only "we calculated overdose mortality rates" without explaining how overdose was defined)	Lower quality of reporting

	Yes – The study provides ICD codes or other information to describe the	Higher quality of reporting
	definition of cause-specific deaths	

Note. ^a Examples of coding for sample representativeness used for guidance comprised:

- 1. A cohort recruited from a single hospital = one sample type, one location
- 2. A cohort recruited through drug treatment and harm reduction services in a single city or neighbourhood = multiple sample types, one location
- 3. A cohort recruited through respondent-driven/street-based/convenience sampling = multiple sample types, multiple locations
- 4. A cohort identified from a database that includes data on people presenting to treatment from a large geographical area = one sample type, multiple locations
- 5. A cohort identified from a database that includes data on people presenting to treatment from a restricted geographical area = one sample type, one location
- 6. A cohort identified from different databases (e.g., drug treatment database, criminal justice database) from a large geographic area = multiple sample types, multiple locations

Exceptions:

- 7. A cohort comprising a very high-risk population for mortality (e.g., people who are incarcerated) = always single sample type and high risk, regardless of number of locations
- 8. A cohort with insufficient information about sample types/site locations = higher risk of bias

Appendix H: Data derived and calculation methods

Crude Mortality Rate (CMR)

If the study failed to report all relevant information related to CMR, the following equations were used to calculate the missing information. Confidence interval and standard error calculations were completed on the log scale to bound the results between 0 and $+\infty$ and to account for deaths being a rarer event.

CMR (per 100 PY) =
$$100 \text{ x} \frac{\text{Observed deaths}}{\text{PYFU}}$$

CMR 95% confidence interval:

$$\ln (CMR \, \text{lci}) = \ln(CMR) - 1.96(\frac{1}{\sqrt{\text{observed deaths}}})$$

$$\ln (CMR uci) = \ln(CMR) + 1.96(\frac{1}{\sqrt{observed deaths}})$$

CMR standard error:

SE (lnCMR) =
$$\frac{1}{\sqrt{\text{observed deaths}}}$$

Where: PY = person years; PYFU = person years of follow-up; lci = lower confidence interval; uci = upper confidence interval; SE = standard error.

Studies that reported deaths and follow up, but not person years

For studies that did not report person years, PYFU were estimated.

The hierarchy of information used to calculate PYFU was 1) drug-specific median FU, 2) drug-specific mean FU, 3) cohort median FU, 4) cohort mean FU, 5) reported FU period then 6) study length.

Due to the proportion of mortality within the cohorts, median was preferred over mean to minimise the potential that the calculated PYFU would be more reflective of the proportion of the cohort who were alive at the conclusion of FU/the study, than the proportion who died. As there were studies who following people who use more than one type of drug, drug-specific data was preferred to better reflective our population of interest; people with regular or problematic use of cocaine. Reported FU period and study length were lower on the hierarchy as there is more potential of bias due to an assumption of when deaths occurred being needed. Reported FU period was given higher preference as the data would be more informative than study length. Where median or mean follow-up period was reported, person year follow up was calculated using:

PYFU = N x (median or mean FU)

For example, if a study reported that people who used cocaine (N = 100) were followed for a median of 2.5 years, PYFU would be calculated as:

PYFU = 100 x 2.5 years = 250 PYFU

For studies where there was no median or mean follow-up period, the study reported follow-up period or study length was used to calculate PYFU using (assuming all deaths occurred halfway through follow-up):

$$PYFU = (n_{alive} x \text{ study period}) + (n_{dead} x \frac{\text{study period}}{2})$$

For example, if a study of 200 people who used cocaine reported that there were 23 deaths within the 6 years of follow-up, PYFU would be calculated as:

PYFU = ((200-23) x 6) + (23 x (6/2)) = (177 x 6) + (23 x 3) = 1131 PYFU

This was our a priori hierarchy for selection of information. However, when we extracted the data, we found no studies that reported median or mean, and thus we were required to use reported FU period and study length.

Standardised Mortality Ratio (SMR)

If the study failed to report all relevant information related to SMR, the following equations were used to calculate the missing information.

$$SMR = \frac{Observed deaths}{Expected deaths}$$

To calculate missing confidence intervals, OpenEpi SMR calculator was used (www.openepi.com/S MR/SMR.htm). OpenEpi is an online, open-source tool created for epidemiologic calculations. The confidence intervals calculated via the Mid-P exact test was used to fill in the missing information. More information can be located on their website about the test (http://www.openepi.com).

SMR standard error:

$$SMR SE = \frac{(SMR uci - SMR lci)}{3.92}$$

Imputing SMRs using data from the Global Burden of Disease study

When CMR was able to be calculated but no relevant SMR data was reported, a SMR was estimated using Global Burden of Disease 2017 (GBD) data. To be able to estimate SMR, all deaths were assumed to occur at the midpoint of the study period. GBD data was accessed through http://ghdx.healthdata.org/gbd-results-tool and matched as closely as possible to the location of study, age range of participants at midpoint and midpoint year of the study. As cocaine use was limited to problematic or harmful use, a prevalence rate for cocaine dependence was calculated using the GBD data as well as a mortality rate for the associated general population. A SMR and relative risk (RR) were then estimated by comparing to study CMR to GBD calculated mortality and prevalence rate.

Relative Risk (RR)

Prior to calculating the RR, the prevalence of cocaine dependence in the population was calculated using Global Burden of Disease 2017 (GBD) data. As with the inputting of SMRs, the data was as closely matched to location of study, age range of participants and year at study midpoint. Estimating RRs from SMRs was based on the methodology outlined by Jones and Swerdlow ⁸.

 $Prevalence in population = \frac{Total number of users}{Total population}$

To calculate RR and its confidence interval, the following equation was used:

Relative Risk =
$$\frac{\text{SMR x (1 - Prevalence in population)}}{1 - (\text{Prevalence in population x SMR)}}$$

Table H1: Contrasting estimated CMRs with those reported in studies with author-reported CMRs

Study	Follow-up (years)	Cohort N	Deaths	PYFU	Reported CMR (95% Cl) (/100 PY)	Calculated PYFU	Calculated CMR (95% CI) (/100PY)
Barrio et al., 2013 ⁹	Mean for cocaine: 5.07	714	9	3922	0.23 (0.12, 0.44)	3619.98 (714 * 5.07)	0.25 (0.13, 0.48)
Callaghan et al., 2013 ¹⁰	01/01/1990 - 31/12/2005 (16)	48949	4356	395738	1.10 (1.07, 1.13)	748336 [((48949 – 4356)*16)+(4356*8)]	0.58 (0.57, 0.60)
Hayashi et al., 2016 ¹¹	Median for all: 5.05	1719	392	11748.63	3.34 (3.02, 3.68)	8680.95 (1719 * 5.05)	4.52 (4.09, 4.99)
Lopez et al., 2004 ¹²	1992-2001 (9)	2212	80	11496	0.70 (0.56, 0.87)	19548 [((2212-80)*9) + (80 *4.5)]	0.41 (0.33, 0.51)
Sanvisens et al., 2014 ¹³	Median for cocaine: 6.5	945	202	7155	2.8 (2.5, 3.2)	6142.5 (945 * 6.5)	3.29 (2.87, 3.78)
van Haastrecht et al., 1996{van Haastrecht, 1996 #217}	01/12/1985 - 01/02/1993 (7.17)	Not reported	9	194	4.64 (2.41, 8.92)	-	-
Vlahov et al., 2008{Vlahov, 2008 #208}	1997 - 31/12/2002 (6)	102	3	485.7	0.62 (0.20, 1.92)	603 [((102- 3)*6)+(3*3)]	0.50 (0.16, 1.54)
Wang et al., 2005{Wang, 2005 #218}	1988 – 31/12/2001 (14)	518	175	3727	4.69 (4.03, 5.45)	6027 [((518- 175)*14)+(175*7)]	2.90 (2.50, 3.36)

Appendix I: Excluded studies

Small cohort (n< 30)

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Case report/case-series

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Not a cocaine using cohort (<90%)

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Cohort defined on the basis of acute care for overdose/poisoning

- Bhattacharya P, Taraman S, Shankar L, Chaturvedi S, Madhavan R. Clinical profiles, complications, and disability in cocaine-related ischemic stroke. Journal of Stroke and Cerebrovascular Diseases. 2011;20(5):443-9. doi:10.1016/j.jstrokecerebrovasdis.2010.02.017
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Subpopulations that have an increased risk of mortality

Severe health presentation or procedure increasing mortality risk

- 1. Kayo N, Shitole S, Beheshtian A, Srinivas V, Scheuer J, Kizer J. Clinical profile, acute care and outcome of cocaine-associated ST-elevation myocardial infarction in a vulnerable urban community. Journal of the American College of Cardiology. 2014;1):A125.
- 2. Shitole SG, Kayo N, Srinivas V, et al. Clinical Profile, Acute Care, and Middle-Term Outcomes of Cocaine-Associated ST-Segment Elevation Myocardial Infarction in an Inner-City Community. *American Journal of Cardiology* 2016; **117**(8): 1224-30.
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Cardiovascular presentations related to cocaine exposure

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- 2. Egbuche O, Ekechukwu I, Maduabum N, Obinwa U, Ukpaka K, Onwuanyi A. Effect of betablocker therapy on readmissions and mortality in heart failure patients with ongoing cocaine use. Journal of Invasive Cardiology 2017; 29 (10): E111.
- 3. Nguyen P, Kamran H, Nasir S, et al. Comparison of Frequency of Cardiovascular Events and Mortality in Patients With Heart Failure Using Versus Not Using Cocaine. Am J Cardiol 2017; 119(12): 2030-4. doi:10.1016/j.amjcard.2017.03.034
- 4. Peralta R, Yoon A, Atoui M, et al. Efficacy of cardiac computerized tomographic angiography in the evaluation of patients with cocaine-induced chest pain: A pilot randomized trial. Circulation Conference: American Heart Association's 2014; 130(Supplement 2): A18655.

HIV positive cohort

- Campa A, Martinez SS, Sherman KE, et al. Cocaine use and liver disease are associated with allcause mortality in the Miami adult studies in HIV (MASH) cohort Journal of Drug Abuse 2016; 2(4). doi:10.21767/2471-853X.100036
- 2. Nacher M, Adenis A, Hanf M, et al. Crack cocaine use increases the incidence of AIDS-defining events in French Guiana. AIDS 2009; 23(16): 2223-6. doi:10.1097/QAD.0b013e32833147c2
- 3. Weber E, Huber M, Battegay M, et al. Influence of noninjecting and injecting drug use on mortality, retention in the cohort, and antiretroviral therapy, in participants in the Swiss HIV Cohort Study HIV Medicine 2015; 16(3): 137-51. doi:10.1111/hiv.12184

Cocaine use not defined as regular or problematic

- 1. Cohen P, Sas A. Ten years of cocaine: A follow-up study of 64 cocaine users in Amsterdam. Amsterdam: Department of Human Geography, University of Amsterdam; 1993.
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- 3. Jones AA, Vila-Rodriguez F, Leonova O, et al. Mortality from treatable illnesses in marginally housed adults: A prospective cohort study. *BMJ Open* 2015; **5 (8) (no pagination)**(e008876): e008876.
- 4. Muhuri PK, Gfroerer JC. Mortality associated with illegal drug use among adults in the United States. *The American Journal of Drug and Alcohol Abuse* 2011; **37**(3): 155-64.

Appendix J: Included studies

Table J1: List of included cohorts (n=21), derived from 28 publications (21 primary and 7 associated secondary publications)

Trials	Papers	Study	Citation	Primary study	Comments	% women	Age at baseline	% PWID	% regular or problematic cocaine use
1	1	Accurso, 2015 ¹⁴	Accurso AJ, Rastegar DA, Ghazarian SR, Fingerhood MI. Impact of hepatitis C status on 20-year mortality of patients with substance use disorders. <i>Addiction Science & Clinical Practice</i> 2015; 10 (20): 1-8. doi:10.1186/s13722-015-0041-6	1					100% dependent on cocaine
2	2	Arendt, 2011 ¹⁵	Arendt M, Munk-Jorgensen P, Sher L, Jensen SOW. Mortality among individuals with cannabis, cocaine, amphetamine, MDMA, and opioid use disorders: A nationwide follow-up study of Danish substance users in treatment. <i>Drug and Alcohol Dependence</i> 2011; 114 (2-3): 134-9.	1					100% dependent on cocaine
3	3	Barrio, 2013 ⁹	Barrio G, Molist G, de la Fuente L, et al. Mortality in a cohort of young primary cocaine users: Controlling the effect of the riskiest drug-use behaviors. <i>Addictive Behaviors</i> 2013; 38 (3): 1601-4. doi:10.1016/j.addbeh.2012.10.007	1					100% report primary drug used as cocaine
4	4	Bohnert, 2017 ¹⁶	Bohnert KM, Ilgen MA, Louzon S, McCarthy JF, Katz IR. Substance use disorders and the risk of suicide mortality among men and women in the US Veterans Health Administration. <i>Addiction</i> 2017; 112 (7): 1193-201. doi:10.1111/add.13774	1		3.87%			100% dependent on cocaine
5	5	Callaghan, 2013 ¹⁰	Callaghan RC, Gatley JM, Veldhuizen S, Lev-Ran S, Mann R, Asbridge M. Alcohol- or drug-use disorders and motor vehicle accident mortality: a retrospective cohort study. <i>Accident; analysis and prevention</i> 2013; 53 : 149-55. doi:10.1016/j.aap.2013.01.008	1		46%			100% dependent on cocaine
	6	Callaghan, 2012 ¹⁷	Callaghan RC, Cunningham JK, Verdichevski M, Sykes J, Jaffer SR, Kish SJ. All-cause mortality among individuals with disorders related to the use of methamphetamine: A comparative cohort study. <i>Drug and Alcohol Dependence</i> 2012; 125 (3): 290-4. doi:10.1016/j.drugalcdep.2012.03.004	0	Secondary paper of Callaghan, 2013 – same recruitment method in similar timeframe resulting in potentially both including the same participants				100% dependent on cocaine
6	7	de la Fuente, 2014 ¹⁸	de la Fuente L, Molist G, Espelt A, et al. Mortality risk factors and excess mortality in a cohort of cocaine users admitted to drug treatment in Spain. <i>Journal of Substance Abuse Treatment</i> 2014; 46 (2): 219-26. doi:10.1016/j.jsat.2013.07.001	1		17.9%	31.0 (mean)	23.1%	100% dependent on cocaine
	8	Brugal, 2016 ¹⁹	Brugal M, Molist G, Sarasa-Renedo A, et al. Assessing gender disparities in excess mortality of heroin or cocaine users compared to the general population. <i>International Journal of Drug Policy</i> 2016; 38 : 36-42. doi:10.1016/j.drugpo.2016.10.009	0	Secondary paper of de la Fuente, 2014 - same recruitment method within similar timeframe resulting in same participants being double counted				100% dependent on cocaine
	9	Molist, 2018 ²⁰	Molist G, Brugal MT, Barrio G, et al. Effect of ageing and time since first heroin and cocaine use on mortality from external and natural causes in a Spanish cohort of drug users. <i>International Journal of Drug Policy</i> 2018; 53 : 8-16. doi:10.1016/j.drugpo.2017.11.011	0	Secondary paper of de la Fuente, 2014 – same recruitment method within similar timeframe resulting in same participants being double counted				100% dependent on cocaine
	10	Colell, 2018 ²¹	Colell E, Domingo-Salvany A, Espelt A, Pares-Badell O, Brugal MT. Differences in mortality in a cohort of cocaine use disorder patients with concurrent alcohol or opiates disorder. <i>Addiction</i> 2018; 113 (6): 1045-55. doi:10.1111/add.14165	0	Secondary paper of de la Fuente, 2014 – same recruitment method within similar timeframe resulting in same participants being double counted				100% dependent on cocaine
7	11	Dias, 2011 ²²	Dias AC, Araujo MR, Dunn J, Sesso RC, de Castro V, Laranjeira R. Mortality rate among crack/cocaine- dependent patients: A 12-year prospective cohort study conducted in Brazil. <i>Journal of Substance</i> <i>Abuse Treatment</i> 2011; 41 (3): 273-8. doi:10.1016/j.jsat.2011.03.008	1		11.45%			100% dependent on cocaine
8	12	Gossop, 2002 ²³	Gossop M, Stewart D, Treacy S, Marsden J. A prospective study of mortality among drug misusers during a 4-year period after seeking treatment. <i>Addiction</i> 2002; 97 (1): 39-47. doi:10.1046/j.1360-0443.2002.00079.x	1					100% report regular use of cocaine
9	13	Hayashi, 2016 ¹¹	Hayashi K, Dong H, Marshall BD, et al. Sex-Based Differences in Rates, Causes, and Predictors of Death Among Injection Drug Users in Vancouver, Canada. <i>American Journal of Epidemiology</i> 2016; 183 (6): 544-52. doi:10.1093/aje/kwv207	1				100%	100% report injecting cocaine in past 6 months

	14	Tyndall, 2001 ²⁴	Tyndall MW, Craib KJP, Currie S, Li K, O'Shaughnessy MV, Schechter MT. Impact of HIV infection of mortality in a cohort of injection drug users. <i>Journal of Acquired Immune Deficiency Syndrome</i> 2001; 28 (4): 351-7. doi:10.1097/00126334-200112010-00008	0	Secondary paper of Hayashi, 2016 – both papers utilised same database of individuals (VIDUS) with overlapping recruitment periods resulting in duplicate data within the 2 papers				100% report injecting cocaine in past 6 months
10	15	Hser, 2012 ²⁵	Hser YI, Kagihara J, Huang D, Evans E, Messina N. Mortality among substance-using mothers in California: a 10-year prospective study. <i>Addiction</i> 2012; 107 (1): 215-22. doi:10.1111/j.1360-0443.2011.03613.x	1		100%			100% dependent on cocaine
11	16	Lopez, 2004 ¹²	Lopez D, Martineau H, Palle C. Mortality of individuals arrested for heroin, cocaine or crack use, 2004.	1					100% arrested for using cocaine
12	17	Markota, 2016 ²⁶	Markota M, Croarkin PE, Bobo WV. Increased 5-year all-cause mortality in youth with positive urine cocaine drug screens. <i>Journal of the American Academy of Child and Adolescent Psychiatry</i> 2016; 55 (10 Supplement 1): S148. doi:10.1016/j.jaac.2016.09.151	1		40%			100% tested positive for cocaine use
13	18	Martell, 2009 ²⁷	Martell BA, Orson FM, Poling J, et al. Cocaine vaccine for the treatment of cocaine dependence in methadone-maintained patients: A randomized, double-blind, placebo-controlled efficacy trial. <i>Archives of General Psychiatry</i> 2009; 66 (10): 1116-23. doi:10.1001/archgenpsychiatry.2009.128	1		33.3%			100% dependent on cocaine
14	19	Nielsen, 2011 ²⁸	Nielsen SF, Hjorthoj CR, Erlangsen A, Nordentoft M. Psychiatric disorders and mortality among people in homeless shelters in Denmark: A nationwide register-based cohort study. <i>The Lancet</i> 2011; 377 (9784): 2205-14. doi:10.1016/S0140-6736(11)60747-2	1		20.95%			100% dependent on cocaine
15	20	O'Driscoll, 2001 ²⁹	O'Driscoll PT, McGough J, Hagan H, Thiede H, Critchlow C, Alexander ER. Predictors of accidental fatal drug overdose among a cohort of injection drug users. <i>American Journal of Public Health</i> 2001; 91 (6): 984-7.	1				71.48%	100% reported primary drug injected consisted of cocaine
16	21	Pavarin, 2017 ³⁰	Pavarin RM, Fioritti A. Mortality Trends among Cocaine Users Treated between 1989 and 2013 in Northern Italy: Results of a Longitudinal Study. <i>Journal of Psychoactive Drugs</i> 2017; 50 (1): 72-80. doi:10.1080/02791072.2017.1365976	1		12.2%	32.8	15.6%	100% dependent on cocaine
	22	Pavarin, 2008 ³¹	Pavarin RM. Cocaine consumption and death risk: A follow-up study on 347 cocaine addicts in the metropolitan area of Bologna. <i>Annali-Istituto Superiore di Sanita</i> 2008; 44 (1): 91-8.	0	Secondary paper of Pavarin, 2017 – same recruitment method in similar timeframes leading to potential overlap of participants between papers				100% dependent on cocaine
	23	Pavarin, 2013 ³²	Pavarin RM. Mortality risk for cocaine abusers in relation to heroin use: A follow-up study. <i>Substance Use & Misuse</i> 2013; 48 (9): 702-10. doi:10.3109/10826084.2013.786731	0	Secondary paper of Pavarin, 2017 – same recruitment method in similar timeframes leading to potential overlap of participants between papers				100% dependent on cocaine
17	24	Ryb, 2009 ³³	Ryb GE, Cooper CC, Dischinger PC, Kufera JA, Auman KM, Soderstrom CA. Suicides, homicides, and unintentional injury deaths after trauma center discharge: Cocaine use as a risk factor. <i>Journal of Trauma - Injury, Infection and Critical Care</i> 2009; 67 (3): 490-6. doi:10.1097/TA.0b013e3181b84430	1		16%			100% tested positive for cocaine use
18	25	Sanvisens, 2014 ¹³	Sanvisens A, Vallecillo G, Bolao F, et al. Temporal trends in the survival of drug and alcohol abusers according to the primary drug of admission to treatment in Spain. <i>Drug and Alcohol Dependence</i> 2014; 136 (1): 115-20. doi:10.1016/j.drugalcdep.2013.12.022	1		26.6%	32	65.8%	100% dependent on cocaine
19	26	van Haastrecht, 1996 ³⁴	van Haastrecht HJA, van Ameijden EJC, van den Hoek JAR, Mientjes GHC, Bax JS, Coutinho RA. Predictors of mortality in the Amsterdam cohort of human immunodeficiency virus (HIV)-positive and HIV-negative drug users. <i>American Journal of Epidemiology</i> 1996; 143 (4): 380-91. doi:10.1093/oxfordjournals.aje.a008752	1		39.56%		77%	100% reported primary drug injected consisted of cocaine
20	27	Vlahov, 2008 ³⁵	Vlahov D, Wang C, Ompad D, et al. Mortality risk among recent-onset injection drug users in five U.S. cities. <i>Substance Use & Misuse</i> 2008; 43 (3-4): 413-28. doi:10.1080/10826080701203013	1				100%	100% reported primary drug injected consisted of cocaine
21	28	Wang, 2005 ³⁶	Wang C, Vlahov D, Galai N, et al. The effect of HIV infection on overdose mortality. <i>AIDS</i> 2005; 19 (9): 935-42. doi:10.1097/01.aids.0000171407.30866.22	1		48.84%		100%	100% reported primary drug injected consisted of cocaine

Table J2. Mortality in included studies

Study (Author, ref)	N	РҮ		All			Drug	-related ^d			Accide	lental injury			S	uicide			н	omicide			Card	iovascular			AID	OS-related		
			N	CMR (95%CI)	SMR (95%CI)	RR (95%Cl)	N	CMR (95%CI)	SMR (95%Cl)	RR (95%CI)	N	CMR (95%CI)	SMR (95%CI)	RR (95%CI)	N	CMR (95%CI)	SMR (95%CI)	RR (95%CI)	N	CMR (95%CI)	SMR (95%CI)	RR (95%Cl)	N	CMR (95%Cl)	SMR (95%Cl)	RR (95%CI)	N	CMR (95%CI)	SMR (95%CI)	RR (95%CI)
Accurso, 2015 ¹⁴	315	5,780*	52	0.90 (0.69- 1.18)	2.11 ^b (1.61- 2.77)	2.12 (1.66- 2.72)																								
Arendt, 2011 ¹⁵	838	2,571.43*	18	0.70 (0.44- 1.11)	6.40 (3.90- 10.00)	6.44 (3.91- 10.09)																								
Barrio, 2013 ⁹	714	3,922	9	0.23 (0.12- 0.44)	4.70 (2.40- 9.00)	4.79 (2.42- 9.40)	3	0.08 (0.02- 0.24)	30.58 ^b (9.86- 94.82)	36.23 (10.35- 187.57)					1	0.03 (0.00- 0.18)	4.19 ^b (0.59- 29.75)	4.26 (4.05- 4.49)									2	0.05 (0.01- 0.20)	47.28 (11.83- 189.07)	
Bohnert, 2017 ¹⁶	83,808														231	0.05 (0.04- 0.06)	3.06 ^b (2.69- 3.48)	3.10 (3.07- 3.11)												
Callaghan, 2013 ¹⁰	48,949	395,738	4,356	1.10 (1.07- 1.13)	2.96ª (2.87- 3.05)	3.00ª (2.90- 3.09)					113	0.03 (0.02- 0.03)	3.80 (2.30- 5.30)	3.87 (2.32- 5.45)																
de la Fuente, 2014 ¹⁸	11,905	65,849.1*	349	0.53 (0.48- 0.59)	4.90 (4.40- 5.40)	4.97 (4.45- 5.49)	117ª	0.18ª (0.17- 0.18)	66.70ª (57.97- 76.75)	86.86ª (72.58- 104.79)	45ª	0.07ª (0.07- 0.07)	8.51 ^a (6.74- 10.50)	8.74 ^a (6.88- 10.86)	34ª	0.05ª (0.05- 0.05)	14.86 ^a (11.21- 19.33)	15.62 ^a (11.63- 20.67)	5ª	0.01ª (0.01- 0.01)	6.42ª (3.43- 10.88)	6.54ª (3.46- 11.27)	45ª	0.07ª (0.07- 0.07)	2.97 ^a (2.35- 3.67)	2.99ª (2.36- 3.70)	4ª	0.02 ^a (0.01- 0.04)	2.85 ^{a,b} (1.07- 7.73)	
Dias, 2011 ²²	131	1,181.62*	27	2.28 (1.57- 3.33)	14.75 (9.92- 21.17)	15.55 (10.26- 22.89)	3	0.25 (0.08- 0.79)	48.79 ^b (15.74- 151.29)	59.39 (16.65- 344.71)	1	0.08 (0.01- 0.60)	1.43 ^b (0.20- 10.17)	1.43 (0.20- 10.53)		•			16	1.35 (0.83- 2.21)	24.39 ^b (14.94- 39.81)	26.72 (15.76- 46.56)					6	0.51 (0.23- 1.13)	38.12 ^b (17.13- 84.86)	
Gossop, 2002 ²³	227						9	0.97 (0.51- 1.87)	141.88 ^b (73.82- 272.67)							94														
Hayashi, 2016 ¹¹	1,719	11,748.6	392	3.34 (3.02- 3.68)	14.05 ^b (12.72- 15.51)	14.64 (10.39- 20.75)	54ª	2.07 (1.19- 3.59)	352.41 ^b (269.91- 460.14)	604.26 ^a (395.99- 1011.87)					7 ª	0.27ª (0.22- 0.33)	15.12 ^{a,b} (7.21- 31.71)	16.77 (13.48- 20.98)					6ª	0.23ª (0.19- 0.28)	0.11 ^{a,b} (0.05- 0.24)	0.11ª (0.05- 0.24)	30ª	1.15 ^a (0.76- 1.73)		
Hser, 2012 ²⁵	511	5,471.48*	27	0.49 (0.34- 0.72)	1.96 ^b (1.35- 2.86)	1.97 (1.35- 2.89)																								
Lopez, 2004 ¹²	2,212	11,496	80	0.70 (0.56- 0.87)	2.09 ^b (1.68- 2.61)	2.10 (1.80- 2.44)																								
Markota, 2016 ²⁶	63	307.5*	3	0.98 (0.31- 3.02)	2.32 ^b (0.75- 7.18)	2.33 (0.77- 7.24)					1	0.33 (0.05- 2.31)	29.49 ^b (4.15- 209.33)		1	0.33 (0.05- 2.31)	18.20 ^b (2.56- 129.22)	20.21 (10.13- 42.69)												
Martell, 2009 ²⁷	115	57.5*	0	-	-	-																								
Nielsen, 2011 ²⁸	525	5,361.5*	75	1.40 (1.12- 1.75)	3.98 ^b (3.17- 4.99)	3.99 (2.90- 5.48)																								
O'Driscoll, 2001 ²⁹	761	2,092.32*	25	1.50 (0.89- 2.54)	5.39 ^b (3.19- 9.10)	5.56 (2.48- 12.79)	30	0.99 (0.69- 1.41)	136.12 ^b (95.17- 194.69)	136.12 (95.68- 193.66)												<u> </u>								
Pavarin, 2017 ³⁰	678	4,752.85*	25	0.53 (0.36- 0.78)	5.21 (3.52- 7.71)	6.30 (4.63- 8.60)	3	0.06 (0.02- 0.20)	25.41 (8.20- 78.78)	26.63 (8.31- 92.19)	6	0.13 (0.06- 0.28)	9.93 (4.46- 22.09)	10.10 (4.49- 23.00)	4	0.08 (0.03- 0.22)	8.33 (3.13- 22.19)	8.45 (3.14- 23.11)	0	-	-	-	0	-	-	-	1	0.02 (0.00- 0.15)	4.98 (0.70- 35.34)	

Study (Author, ref)	N	PY		All	deaths			Drug	-related ^d			Accide	ntal injury			S	uicide			H	omicide			Card	iovascular			AID	S-related	
Ryb, 2009 ³³	2,451	15,931.5*									32	0.20 (0.14- 0.28)	5.72 ^b (4.04- 8.08)	5.92 (4.13- 8.51)	5	0.03 (0.01- 0.08)	1.86 ^b (0.77- 4.46)	1.87 (1.82- 1.92)	12	0.08 (0.04- 0.13)	5.13 ^b (2.91- 9.04)	5.29 (2.95- 9.58)								
Sanvisens, 2014 ¹³	2,451	7,155	202	2.80 (2.50- 3.20)	6.12 ^b (5.47- 7.00)	6.21 (5.53- 7.11)																	9	0.13 (0.10- 0.24)	16.29 ^b (8.48- 31.31)	16.96 (8.64- 33.97)	59	0.82 (0.64- 1.06)	30.28 ^b (23.46- 39.08)	
van Haastrecht, 1996 ³⁴		194	9	4.64 (2.41- 8.92)	84.93 ^b (44.19- 163.24)																									
Vlahov, 2008 ³⁵	102	485.7	3	0.62 (0.20- 1.92)	5.48 ^b (1.77- 17.00)	5.72 (2.77- 12.15)																								
Wang, 2005 ³⁶	518	3727	175	4.69 (4.03- 5.45)	13.02 ^b (11.19- 15.13)	14.09 (11.96- 16.61)	29	0.78 (0.54- 1.12)	100.12 ^b (69.57- 144.07)	267.11 (142.16- 790.65)					2	0.05 (0.01- 0.21)	3.15 ^b (0.79- 12.58)	3.19 (2.96- 3.44)					7	0.19 (0.09- 0.39)	1.95 (0.93- 4.09)	1.96 (0.93- 4.17)	62	1.66 (1.30- 2.13)	63.87 ^b (49.79- 81.92)	

* PYFU was calculated using formula in Appendix H; * Information from secondary paper was used; * SMR was inputted using GBD 2017 estimates (see Appendix H for more information); c Insufficient data was reported to estimate confidence intervals. * Drug-related': note that estimates could include poisoning deaths, deaths attributed to mental and behavioural disorders due to psychoactive substance use, and deaths from other causes that were thought to be drug-related.

Table J3. ICD codes used to identify each cause-specific death

Study (Author,ref)	Drug-related	Suicide	Accidental Injury	Cardiovascular disease	Homicide	AIDS-related	Cancer	Respiratory disease
Barrio, 2013 ⁹	-	-	-	-	-	-	-	-
Bohnert, 2017 ¹⁶	-	ICD 10: X60-84, Y87.0	-	-	-	-	-	-
Brugal, 2016 ^{19a}	ICD 9: 291, 292, 303-305, E850-869, E929.2, E980, 427.5 ^b , 514 ^b , 518.4 ^b , 780-799 ^b , 980 ^b ; ICD 10: F10-F19, F55, X40-X49, Y10-Y19, I46 ^b , J81 ^b , J96 ^b , R00-R74 ^b , R76-R99 ^b	-	-	-	-	-	-	-
Callaghan, 2013 ¹⁰	-	-	ICD-9 810-825 (with a 4th digit of "0" or "2"); ICD-10 V20-29 (3rd digit of "0" or "4"), V30-79 (3rd digit of "0" or "5"), and V83-86 (3rd digit of "0" or "5")	-	-	-	-	-
Colell, 2018 ^{21a}	-	-	Re	ICD 10: J81, J96, J960, J969, J13, J18, J189, 496, J09, J209, J439, J441, J449, J841, J850, J90, J988	-	ICD 9: 279.5; ICD10: B20, B200, B202, B203, B205, B206, B207, B208, B209, B21, B212, B213, B218, B220, B222, B227, B238, B24	-	-
Dias, 2011 ²²	-	-	-	-	-	-	-	-
Gossop, 2002 ²³	-	-	-	-	-	-	-	-
Markota, 2016 ²⁶	-	-	-	-		-	-	-
Molist, 2018 ^{20a}	ICD 9: 291, 292, 303-305, 427.5 ^b , 514 ^b , 518.4 ^b , 780- 799 ^b , 980.0 ^b , E850-869, E929.2, E980; ICD 10: F10- F19, F55, I46 ^b , J81 ^b , J96 ^b , R00-R74 ^b , R76-R99 ^b , X40- X49, Y10-Y19, Y90-Y91	ICD 9: E950-E959; ICD 10: X60-X84, Y87.0	ICD 9: E800-E949, E980- E989; ICD 10: V01-X59, Y10-Y34, Y40-Y86, Y87.2, Y88, Y89.9	ICD 9: 390-459; ICD 10: 100-199	ICD 9: E960-E979, E990-E999; ICD 10: X85-Y09, Y35-Y36, Y87.1, Y89.0-Y89.1		ICD 9: 140-239; ICD 10: C00-D49	
O'Driscoll, 200129	-	-	-	-	-	-	-	-
Pavarin, 2017 ³⁰	-	-	-	-	-	-	-	-
Ryb, 2009 ³³	-	-	-	-	-	-	-	-
Sanvisens, 2014 ¹³	-	-	-	-	-	ICD-9: 279.5, 795.8; ICD- 10: B20–B24, R75	ICD-9: 140–154, 156– 208, 273.3; ICD10: C00–C21, C23–C97, D00–D09, D37–D48	-
Tyndall, 2001 ^{24c}	-	-	-	-	-	-	-	-
Wang, 2005 ³⁶	-	-	-	-	-	-	-	-

Appendix K: Relative risk results

Study		%
ID	RR (95% CI)	Weight
High-Income North America		
Accurso, 2015 (United States)	2.12 (1.61, 2.80)	7.11
Markota, 2016 (United States)	2.33 (0.75, 7.45)	4.51
Callaghan, 2012 (United States)	3.00 (2.90, 3.09)	7.36
O'Driscoll, 2001 (United States)	5.56 (3.24, 9.63)	6.45
Vlahov, 2008 (United States)	5.72 (1.78, 19.92)	4.34
Wang, 2005 (United States)	14.09 (11.96, 16.61)	7.27
Hayashi, 2016 (Canada)	➡ 14.64 (13.21, 16.25)	7.33
Subtotal (I-squared = 99.5%, p = 0.000)	5.30 (2.37, 11.83)	44.39
• • • • • • • • • • • •	T	
Western Europe	1	
Lopez, 2004 (France)	2.10 (1.68, 2.61)	7.20
Nielsen, 2011 (Denmark)	3.99 (3.18, 5.01)	7.19
Barrio, 2013 (Spain)	4.79 (2.42, 9.40)	6.04
de la Fuente, 2014 (Spain)	4.97 (4.45, 5.49)	7.33
Sanvisens, 2014 (Spain)	← 6.21 (5.53, 7.11)	7.31
Pavarin, 2017 (Italy)	6.30 (4.63, 8.60)	7.05
Arendt, 2011 (Denmark)	6.44 (3.91, 10.09)	6.65
Subtotal (I-squared = 92.3%, p = 0.000)	4.63 (3.47, 6.17)	48.77
Tropical Latin America	1	
Dias, 2011 (Brazil)	15.55 (10.26, 22.89)	6.84
Subtotal (I-squared = .%, p = .)	15.55 (10.41, 23.22)	6.84
Overall (I-squared = 98.9%, p = 0.000)	5.41 (3.65, 8.01)	100.00
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Figure K1. Pooled estimates, derived from random effects meta-analysis, of all-cause mortality relative risk among people with regular or problematic use of cocaine, overall and by region.

Study ID	RR (95% CI)	% Weight
Drug Poisonings	CONTRACTOR OF THE STATE	
Barrio, 2013 (Spain)	36.23 (10.35, 187.57)	10.77
Dias, 2011 (Brazil)	59.39 (16.65, 344.71)	10.35
Molist, 2018 (Spain)	 86.86 (72.58, 104.79) 	18.69
O'Driscoll, 2001 (United States)	136.12 (95.17, 194.69)	18.08
Pavarin, 2017 (Italy)	26.63 (8.31, 92.19)	12.43
Tyndall, 2001 (Canada)	604.26 (395.99, 1011.87)	17.53
Wang, 2005 (United States)	267.11 (122.60, 1476.51)	12.14
Subtotal (I-squared = 91.2%, p = 0.000)	114.60 (55.54, 236.47)	100.00
Accidental Injury		
Callaghan, 2013 (United States)	3.87 (2.32, 5.45)	24.65
Dias, 2011 (Brazil)	1.43 (0.20, 10.53)	3.99
Molist, 2018 (Spain)	8.74 (6.88, 10.86)	30.08
Pavarin, 2017 (Italy)	10.10 (4.49, 23.00)	14.74
Ryb, 2009 (United States)	5.92 (4.13, 8.51)	26.55
Subtotal (I-squared = 73.5%, p = 0.005)	6.13 (4.02, 9.34)	100.00
Suicide		
Barrio, 2013 (Spain)	4.26 (0.59, 35.06)	8.19
Bohnert, 2017 (United States)	3.10 (2.71, 3.53)	16.64
Markota, 2016 (United States)	20.21 (2.59, 501.21)	6.13
Molist, 2018 (Spain)	15.62 (11.63, 20.67)	16.38
Pavarin, 2017 (Italy)	8.45 (3.14, 23.11)	13.39
Ryb, 2009 (United States)	1.87 (0.77, 4.57)	13.96
Tyndall, 2001 (Canada)	16.77 (7.55, 40.19)	14.23
Wang, 2005 (United States)	3.19 (0.79, 13.65)	11.08
Subtotal (I-squared = 94.1%, p = 0.000)	6.32 (2.79, 14.34)	100.00
Cardiovascular Disease		
Molist, 2018 (Spain)	2.99 (2.36, 3.70)	25.99
Sanvisens, 2014 (Spain)	16.96 (8.64, 33.97)	24.89
Tyndall, 2001 (Canada)	0.11 (0.05, 0.24)	24.46
Wang, 2005 (United States)	1.96 (0.93, 4.17)	24.65
Subtotal (I-squared = 96.7%, p = 0.000)	1.85 (0.39, 8.85)	100.00
Homicide		00.75
Dias, 2011 (Brazil)	26.72 (15.76, 46.56)	33.72
Molist, 2018 (Spain)	6.54 (3.46, 11.27)	33.13
Ryb, 2009 (United States)	5.29 (2.95, 9.58)	33.15
Subtotal (I-squared = 89.5%, p = 0.000)	9.80 (3.52, 27.25)	100.00
I I I .01 1 10	100 2000	

Figure K2. Forest plots of cause-specific mortality relative risk among people with regular or problematic use of cocaine.

	Pooled mortality relative risk (95%CI)	 ²	References
All-cause mortality	5.41 (3.65-8.01)	98.9%	9,11-15,17,18,22,26,28-30,35,36
Women	5.24 (3.25-8.44)	95.2%	15,17,18,25,28,31
Men	3.44 (2.86-4.15)	74.3%	15,17,18,26,28,31
GBD Region			
High-income North America	5.35 (2.80-10.25)	98.4%	11,14,17,26,29,35,36
Western Europe	4.64 (3.24-6.63)	95.4%	9,12,13,15,18,28,30
Tropical Latin America	15.55 (10.26-22.89)	-	22
Cause-specific mortality			
Drug-related ^a	114.60 (55.54-236.47)	91.2%	9,20,22,24,29,30,36
Accidental injury	6.13 (4.02-9.34)	73.5%	10,20,22,26,30,33
Suicide	6.32 (2.79-14.34)	94.1%	9,16,20,24,26,30,33,36
Cardiovascular disease	1.85 (0.39-8.85)	96.7%	13,20,24,36
Homicide	9.80 (3.52-27.25)	89.5%	20,22,33
AIDS-related	13.54 (3.78-48.56)	88.2%	13,21,22,30
Cancer	1.49 (0.70-3.18)	85.2%	13,20,24,30,36
Liver-related	3.39 (0.50-23.02)	92.5%	13,21,22
Cerebrovascular disease		-	-
Digestive disease	-	-	-
Nervous system		-	-
Respiratory disease	-	-	-

Table K1. Summary of pooled estimated relative risks of mortality among people with regular or problematic use of cocaine

Note. ^a 'Drug-related': note that estimates could include poisoning deaths, deaths attributed to mental and behavioural disorders due to psychoactive substance use, and deaths from other causes that were thought to be drug-related.

Appendix L: Supplementary results

	No. studies	No. people	Pooled crude mortality rate (95%Cl)	l ²	References
All-cause mortality	10	67,159	1.25 (0.81-1.94)	99.1%	9-13,15,18,22,30,34-36
Women ^a	4	24,596 ^b	0.67 (0.56-0.81)	44.6%	10,15,18,32
Men	4	36,588 ^b	0.70 (0.33-1.45)	99.0%	10,15,18,32
GBD Region					
High-income North America	4	51,288	1.98 (0.83-4.71)	99.6%	10,11,35,36
Western Europe	7	17,292 ^c	0.88 (0.42, 1.85)	98.7%	9,12,13,15,18,30,34 t
Tropical Latin America	1	131	2.29 (1.57-3.33)	-	22
Cause-specific mortality					
Drug-related ^d	-3	12,750	0.10 (0.05-0.19)	52.5%	9,19,22
Accidental injury	3	60,985	0.05 (0.02-0.11)	92.1%	10,20,22
Suicide	2	12,619	0.05 (0.04-0.07)	0.0%	9,20
Cardiovascular disease	2	12,850	0.09 (0.05-0.15)	63.2%	13,20
Homicide	2	12,036	0.10 (0.00-16.41)	99.0%	20,22
AIDS-related	4	5,380	0.14 (0.02-0.88)	95.7%	9,13,21,22
Respiratory disease	2	4,304	0.06 (0.04-0.09)	0.0%	9,21
Cancer	2	12,850	0.06 (0.05-0.08)	0.0%	13,20
Digestive disease	1	11,905	0.03 (0.02-0.05)	-	20

Table L1. Summary of pooled results using only study-reported CMRs (i.e. not including imputedCMR estimates) among people with regular or problematic use of cocaine

^a Pavarin, 2013 included, but no women died. ^b Except for Arendt, 2011, all studies reported the N of cocaineusing participants ^c Except for van Haastrecht, 1996, all studies reported the N of cocaine-using participants. ^d 'Drug related': note that estimates could include poisoning deaths, deaths attributed to mental and behavioural disorders due to psychoactive substance use, and deaths from other causes that were thought to be drug-related.

			Pooled		
	No. studies	No. people	standardised mortality ratio (95%Cl)	²	References
All-cause mortality	7	63,431	5.58 (3.90-7.99)	96.6%	9,15,17,18,22,30,37
Women	3	24,538ª	7.00 (3.12-15.68)	96.7%	15,17,18
Men	4	36,603ª	3.41 (2.74-4.23)	84.3%	15,17,18,31
GBD Region					
High-income North America	2	49,165	3.94 (2.11-7.35)	89.3%	17,37
Western Europe	4	14,135	4.97 (4.51-5.47)	0.0%	9,15,18,30
Tropical Latin America	1	131	14.75 (10.10-21.55)	-	22
Cause-specific mortality					
Drug-related ^b	2	12,583	44.37 (37.28-52.81)	0.0%	20,30
Accidental injury	3	61,532	5.01 (3.71-6.77)	29.9%	10,20,30
Suicide	2	12,583	9.25 (6.73-12.73)	0.0%	20,30
Cardiovascular disease	2	12,583	2.62 (1.21-5.68)	66.8%	20,30
Homicide	1	11,905	4.17 (1.69-10.25)	-	20
AIDS-related	1	678	4.98 (0.70-35.34)	-	30
Respiratory disease	1	678	24.12 (6.03-96.43)	-	30
Cancer	2	12,583	1.10 (0.61-1.98)]	20,30
Digestive disease	1	11,905	1.90 (1.20-2.89)	-	20

Table L2. Summary of pooled results using only study-reported SMRs (i.e. not including imputedSMR estimates) among people with regular or problematic use of cocaine

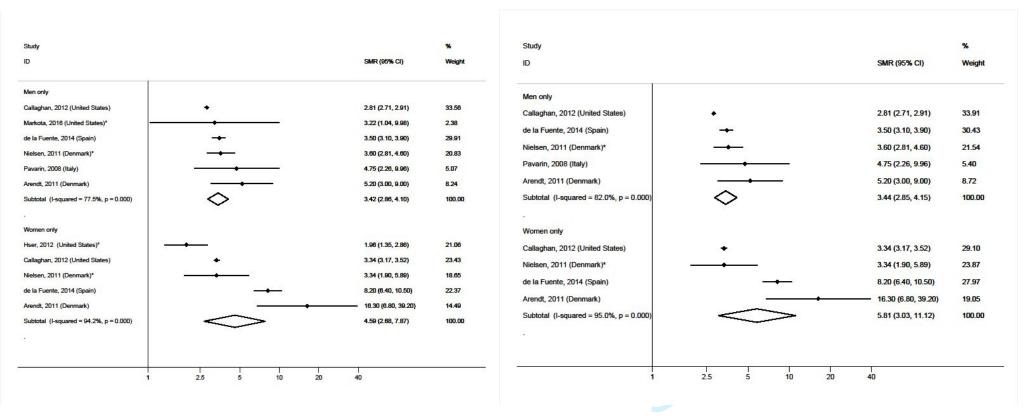
^a Except for Arendt, 2011, all studies reported the N of cocaine-using participants. ^b 'Drug related': note that estimates could include poisoning deaths, deaths attributed to mental and behavioural disorders due to psychoactive substance use, and deaths from other causes that were thought to be drug-related.

Figure L1. Pooled estimates, derived from random effects meta-analysis, for all-cause CMR by gender (men only vs. women only) per 100 person-years with studies that reported deaths by either gender (left) or both genders (right).

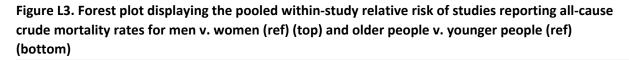
Study		%	Study		%
D	CMR/100PY (95% CI)	Weight	ID	CMR/100PY (95% CI)	Weight
len only			Men only		
Pavarin, 2013 (Italy)	0.39 (0.18, 0.87)	14.15	Pavarin, 2013 (Italy)	0.39 (0.18, 0.87)	15.93
le la Fuente, 2014 (Spain) -	0.53 (0.47, 0.59)	19.54	de la Fuente, 2014 (Spain) -		21.95
Arendt, 2011 (Denmark)	0.66 (0.38, 1.14)	16.62	Arendt, 2011 (Denmark)	0.66 (0.38, 1.14)	18.69
Callaghan, 2013 (United States) +	1.47 (1.42, 1.53)	19.68	Callaghan, 2013 (United States)	 1.47 (1.42, 1.53) 	22.10
Nielsen, 2011 (Denmark)	1.49 (1.17, 1.91)	18.99	Nielsen, 2011 (Denmark)		21.33
Markota, 2016 (United States)	1.64 (0.53, 5.10)	11.02			
Subtotal (I-squared = 98.3%, p = 0.000)	0.89 (0.50, 1.56)	100.00	Subtotal (I-squared = 98.7%, p = 0.000)	0.82 (0.45, 1.50)	100.00
8			-		
Nomen only			Women only		
Hser, 2012 (United States)	0.49 (0.34, 0.72)	17.54	de la Fuente, 2014 (Spain)	• 0.56 (0.42, 0.70)	29.14
de la Fuente, 2014 (Spain)	0.56 (0.42, 0.70)	25.80	Callaghan, 2013 (United States)	 ◆ 0.72 (0.68, 0.76) 	56.62
Callaghan, 2013 (United States)	0.72 (0.68, 0.76)	41.73	Arendt, 2011 (Denmark)	0.82 (0.34, 1.98)	4.47
Arendt, 2011 (Denmark)	0.82 (0.34, 1.98)	4.83	Nielsen, 2011 (Denmark)	1.05 (0.60, 1.85)	9.78
Nielsen, 2011 (Denmark)	- 1.05 (0.60, 1.85)	10.09	Subtotal (I-squared = 44.5%, p = 0.145)	0.70 (0.57, 0.85)	100.00
Subtotal (I-squared = 55.6%, p = 0.061)	0.66 (0.54 <mark>,</mark> 0.81)	100.00	Contract (Fordation 11.000, p. 21.10)		
Α					

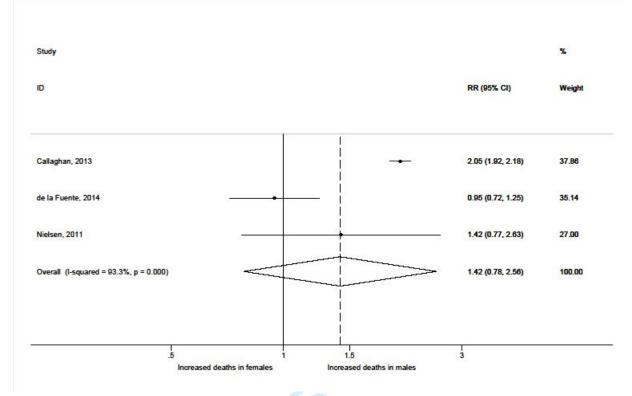
Note: Pavarin, 2013³² reported no deaths among women and therefore not included in pooled estimate.

Figure L2. Pooled estimates, derived from random effects meta-analysis, for all-cause SMR by gender (men only vs. women only) per 100 person-years with studies that reported deaths by either gender (left) or both genders (right).



Note: Pavarin, 2008³¹ reported no deaths among women and therefore not included in pooled estimate.





Note: Arendt, 2011¹⁵ was not included in the gender within-study relative risk analysis as they did not report number of deaths within men and women

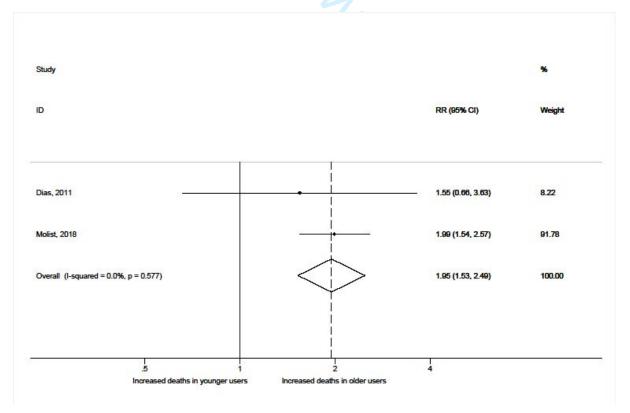


Figure L4. Forest plots derived from random effects meta-analysis for younger people (< 30) vs. older people (≥ 30) for all-cause CMR per 100 person-years (left) and all-cause SMR (right).

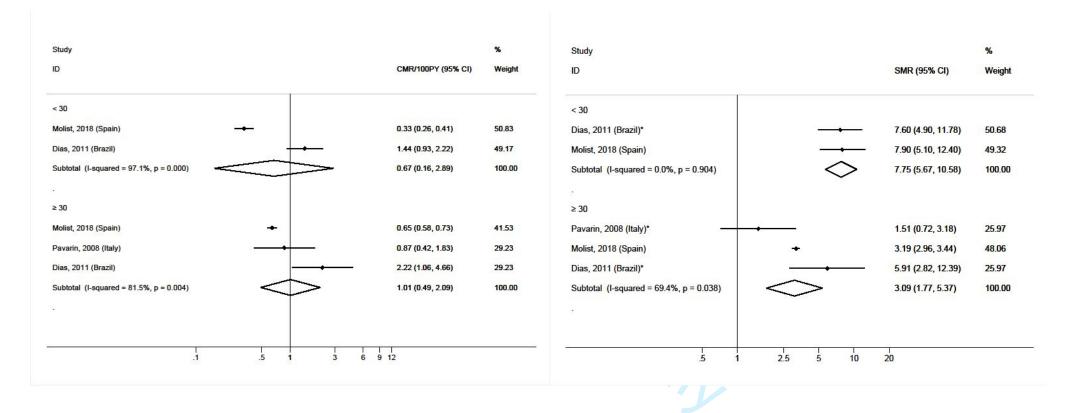


Figure L5. Forest plots derived from random effects meta-analysis by sampling frame: national vs. subnational (including states/provinces) vs. city for all-cause CMR per 100 person-years (left) and all-cause SMR (right).

udy	CMR/100PY (95% CI)	% Weight	Study ID		SMR (95% CI)	% Weigh
itional			National			
ahov, 2008 (United States)	0.62 (0.20, 1.92)	4.18	Lopez, 2004 (France)*	i i i	2.09 (1.68, 2.61)	6.75
pez, 2004 (France)	0.70 (0.56, 0.87)	6.80	Nielsen, 2011 (Denmark)*	i	3.98 (3.17, 4.99)	6.74
endt, 2011 (Denmark)	0.70 (0.44, 1.11)	6.27	Vlahov, 2008 (United States)*		5.48 (1.77, 17.00)	4.36
elsen, 2011 (Denmark)		6.79	Arendt, 2011 (Denmark)		6.40 (3.90, 10.00)	6.26
ibtotal (I-squared = 85.8%, p = 0.000)	0.86 (0.53, 1.37)	24.04	Subtotal (I-squared = 88.9% , p = 0.000)	\sim	3.85 (2.26, 6.56)	24.11
				\sim	5.55 (2.20, 0.50)	
bnational			Subnational			
nrio, 2013 (Spain)	0.23 (0.12, 0.44)	5.71	Markota, 2016 (United States)*		2.32 (0.75, 7.18)	4.36
la Fuente, 2014 (Spain)	0.53 (0.48, 0.59)	6.93	Callaghan, 2012 (United States)		2.96 (2.87, 3.05)	6.89
arkota, 2016 (United States)	0.98 (0.31, 3.02)	4.18	Barrio, 2013 (Spain)		4.70 (2.40, 9.00)	5.75
Illaghan, 2013 (United States)	• 1.10 (1.07, 1.13)	6.96	de la Fuente, 2014 (Spain)	I	4.90 (4.40, 5.40)	6.86
Driscoll, 2001 (United States)	1.50 (0.89, 2.54)	6.10	O'Driscoll, 2001 (United States)*		5.39 (3.19, 9.10)	6.13
btotal (I-squared = 98.0%, p = 0.000)	0.73 (0.43, 1.22)	29.89	Subtotal (I-squared = 95.6%, p = 0.000)	Ň.	4.03 (2.78, 5.84)	29.99
ty .				i		
varin, 2017 (Italy)	0.53 (0.36, 0.78)	6.45	City	i i	2 44 /4 64 2 77)	6.67
curso, 2015 (United States)	◆ 0.90 (0.69, 1.18)	6.71	Accurso, 2015 (United States)*		2.11 (1.61, 2.77)	
as, 2011 (Brazil)	2.29 (1.57, 3.33)	6.49	Pavarin, 2017 (Italy)		5.21 (3.52, 7.71)	6.44
invisens, 2014 (Spain)	2.29 (1.57, 3.53) → 2.80 (2.50, 3.20)	6.91	Sanvisens, 2014 (Spain)*	T.	6.12 (5.47, 7.00)	6.85
invisens, 2014 (Spain) iyashi, 2016 (Canada)	→ 2.00 (2.50, 3.20)	6.93	Wang, 2005 (United States)*		13.02 (11.19, 15.13)	6.82
n Haastrecht, 1996 (Netherlands)	4.64 (2.41, 8.92)	5.70	Hayashi, 2016 (Canada)*	I •	14.05 (12.72, 15.51)	6.86
ang, 2005 (United States)		6.89	Dias, 2011 (Brazil)	_	14.75 (9.92, 21.17)	6.47
	4.69 (4.03, 5.45)	46.08	van Haastrecht, 1996 (Netherlands)*		44.19 , 163.24)	5.77
btotal (I-squared = 96.9%, p = 0.000)	2.16 (1.42, 3.27)	40.08	Subtotal (I-squared = 98.0%, p = 0.000)		10.14 (6.08, 16.91)	45.90
	1.24 (0.86, 1.78)	100.00	Overall (I-squared = 99.0%, p = 0.000)		6.13 (4.15, 9.05)	100.0

Figure L6. Forest plots derived from random effects meta-analysis by cocaine form used: cocaine vs. cocaine and crack cocaine vs. speedball vs. cocaine and heroin for all-cause CMR per 100 person-years (left) and all-cause SMR (right).

0.23 (0.12, 0.44) 0.53 (0.36, 0.78)	5.99	ID Cocaine		SMR (95% CI)	Weight
0.53 (0.36, 0.78)	5.99	Contraction of the second se			
0.53 (0.36, 0.78)	5.99				
	6.75	Muhuri, 2011 (United States)*		1.08 (0.88, 1.28)	5.67
0.53 (0.48, 0.59)	7.24	Hser, 2012 (United States)*		1.96 (1.35, 2.86)	5.41
0.69 (0.49, 0.93)	6.92	Accurso, 2015 (United States)*	—	2.11 (1.61, 2.77)	5.57
0.70 (0.44, 1.11)	6.56	Cohen, 1993 (Netherlands)*		2.23 (0.72, 6.92)	3.68
		Markota, 2016 (United States)*		2.32 (0.75, 7.18)	3.68
		Callaghan, 2012 (United States)	•	2.96 (2.87, 3.05)	5.75
		Hakansson 2013 (Sweden)*	_ _	3 09 (2 11 4 55)	5.40
1.40 (1.12, 1.75)	7.09				5.63
	6.38				4.83
2.29 (1.57, 3.33)	6.78		1		5.73
 2.80 (2.50, 3.20)	7.22				5.39
 3.34 (3.02, 3.68) 	7.24				5.13
4.64 (2.41, 8.92)	5.97				5.51
4.69 (4.03, 5.45)	7.19				5.71
1.27 (0.87, 1.86)	100.00				5.24
					5.69
			•		5.73
		the second se		the second s	5.41
					4.84
		Subtotal (I-squared = 98.9%, p = 0.000)	\diamond	5.17 (3.60, 7.43)	100.00
2.38 (1.35, 4.18)	100.00				
					10.05
0 70 (0 56 0 87)	100 00		—		43.65
			•		56.35
	1.1.1.1.1.1.1	Subtotal (I-squared = 87.0%, p = 0.006)	\sim	13.42 (6.81, 26.44)	100.0
0.77 (0.57, 1.04)	79.25				
0.95 (0.52, 1.71)	20.75		-	2.09 (1.68, 2.61)	100.00
0.80 (0.61, 1.05)	100.00	Subtotal (I-squared = .%, p = .)	•	2.09 (1.68, 2.61)	100.00
	0.90 (0.69, 1.18) 0.98 (0.31, 3.02) 1.10 (1.07, 1.13) 1.40 (1.12, 1.75) 1.50 (0.89, 2.54) 2.29 (1.57, 3.33) ↓ 3.34 (3.02, 3.68) ↓ 4.69 (4.03, 5.45) 1.27 (0.87, 1.86) ↓ 0.94 (0.57, 1.53) 1.99 (1.88, 2.10) 3.51 (2.39, 5.15) 3.51 (2.39, 5.15) 4.37 (4.07, 4.68) 2.38 (1.35, 4.18) 0.70 (0.56, 0.87) 0.70 (0.56, 0.87) 0.77 (0.57, 1.04) 0.95 (0.52, 1.71)	0.90 (0.69, 1.18) 0.98 (0.31, 3.02) 4.40 1.10 (1.07, 1.13) 7.27 1.40 (1.12, 1.75) 7.09 1.50 (0.89, 2.54) 4.63 2.29 (1.57, 3.33) 4.37 (3.02, 3.68) 7.24 4.64 (2.41, 8.92) 5.97 4.69 (4.03, 5.45) 7.19 1.27 (0.87, 1.86) 100.00 0.94 (0.57, 1.53) 22.38 1.99 (1.88, 2.10) 26.87 3.51 (2.39, 5.15) 23.93 4.37 (4.07, 4.68) 2.38 (1.35, 4.18) 100.00 0.70 (0.56, 0.87) 100.00 0.77 (0.57, 1.04) 7.925 0.95 (0.52, 1.71) 20.75	0.90 (0.69, 1.18) 0.90 (0.69, 1.18) 0.98 (0.31, 3.02) 4.40 Callaghan, 2012 (United States) 1.10 (1.07, 1.13) 7.27 Hakansson, 2013 (Sweden)* 1.40 (1.12, 1.75) 7.09 Nielsen, 2011 (Denmark)* 2.29 (1.57, 3.33) 6.78 de la Fuente, 2014 (Spain) 2.29 (1.57, 3.33) 6.78 de la Fuente, 2014 (Spain) 2.29 (1.57, 3.33) 6.78 de la Fuente, 2014 (Spain) 4.69 (4.03, 5.45) 7.19 Vlahov, 2008 (United States)* 4.69 (4.07, 1.53) 22.38 Hayashi, 2016 (Canada)* Uas, 2011 (Brazil) 4.37 (4.07, 4.68) 26.82 Subtotal (I-squared = 98.9%, p = 0.000) 2.38 (1.35, 4.18) 100.00 Cocaine + Heroin Pavarin, 2017 (Italy) 0.70 (0.56, 0.87) 100.00 0.77 (0.57, 1.04) 79.25 Cocaine + Crack cocaine Lopez, 2004 (France)*	0.90 (0.54, 1.17) 0.30 0.90 (0.59, 1.18) 7.01 0.98 (0.31, 3.02) 4.40 1.10 (1.07, 1.13) 7.27 Hakansson, 2013 (Sweden)* + 1.40 (1.12, 1.75) 7.09 Nielsen, 2011 (Denmark)* + 229 (1.57, 3.33) 6.78 229 (1.57, 3.33) 6.78 280 (2.50, 3.20) 7.22 Pavarin, 2017 (Italy) + 4.64 (2.41, 8.92) 5.97 Vlahov, 2008 (United States)* + 4.64 (2.41, 8.92) 5.97 Vlahov, 2008 (United States)* + 4.69 (4.03, 5.45) 7.19 Sanvisens, 2014 (Spain)* + 4.69 (4.03, 5.45) 7.19 Sanvisens, 2014 (Spain)* + 4.69 (4.03, 5.45) 7.19 Sanvisens, 2014 (Spain)* + 1.99 (1.88, 2.10) 26.87 Disa, 2011 (Brazi)* + 4.37 (4.07, 4.68) 26.82 Subtotal (I-squared = 98.9%, p = 0.000) + 2.38 (1.35, 4.18) 100.00 0.70 (0.56, 0.87) 100.00 0.70 (0.56	$\begin{array}{c cccc} 0.50 & (0.647, 1.18) & 7.01 & 0.30 & 0.68, 1.18) & 7.01 & 0.98 & (0.31, 3.02) & 4.40 & Callaghan, 2012 (United States) & 2.32 (0.75, 7.18) \\ 0.98 & (0.31, 3.02) & 4.40 & Callaghan, 2012 (United States) & 2.96 (2.87, 3.05) \\ 1.10 & (1.07, 1.13) & 7.27 & Hakansson, 2013 (Sweden)^* & 3.98 (2.11, 4.55) \\ 1.40 & (1.12, 1.75) & 7.09 & Nielsen, 2011 (Denmark)^* & 3.98 (2.11, 4.55) \\ 1.50 & (0.89, 2.54) & 6.38 & Barrio, 2013 (Spain) & & & & & & & & & & & & & & & & & & &$

Figure L7. Bubble plot displaying results of meta-regression analysis examining the impact of proportion of women within cohort on CMR per 100PY.

Meta-regression	1				Number of obs	=	9
REML estimate of	of between-s	tudy variance	e		tau2	=	.9327
% residual var:	iation due to	o heterogene:	ity		I-squared_res	=	99.11%
Proportion of h	between-study	y variance e	xplained		Adj R-squared	=	-9.76%
With Knapp-Hart	L estimate of between-study variance tau2 = .932 esidual variation due to heterogeneity I-squared_res = 99.11 portion of between-study variance explained Adj R-squared = -9.76 h Knapp-Hartung modification ln_cmr exp(b) Std. Err. t P> t [95% Conf. Interval						
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
prop_women	.1940122	.5429045	-0.59	0.576	.0002595	1	45.0524
cons	1.717208	1.303015	0.71	0.499	.2854824	1	0.32919

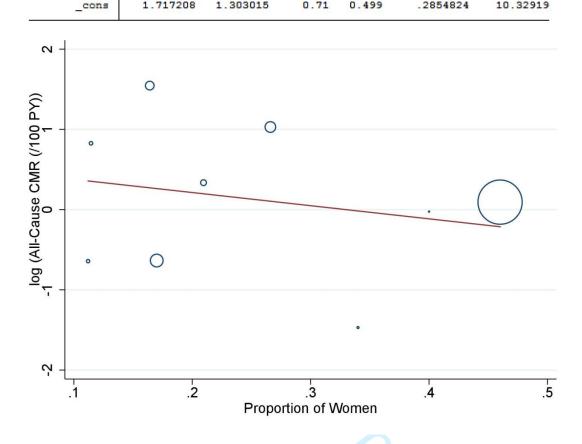


Figure L8. Bubble plot displaying results of meta-regression analysis examining the impact of proportion of people who inject drugs within cohort on CMR per 100PY.

Meta-regression	n				Number of obs	=	8
REML estimate	of between-s	tudy varianc	e		tau2	=	.2667
% residual var.	iation due to	o heterogene.	ity		I-squared res	=	92.04%
Proportion of 1	between-stud	y variance e	xplained		Adj R-squared	=	68.62%
With Knapp-Har	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
inject	5.870798	3.271986	3.18	0.019	1.501173	2	2.95955
cons	.5090705	.2276331	-1.51	0.182	.1704495	1	. 520408

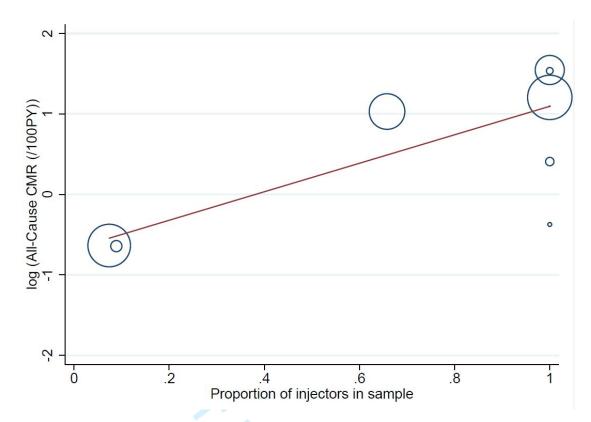


Figure L9. Meta-regression analysis examining the impact of cocaine form used (ref. Cocaine/Cocaine and Crack Cocaine; Cocaine 1 = Speedball/Cocaine and Heroin) on CMR per 100PY.

Meta-regression			-	Num	ber of obs	= 22
REML estimate o:	f between-st	udy variance		tau	2	= .6419
<pre>% residual varia</pre>	ation due to	heterogeneit	су	I-s	quared_res	= 98.82%
Proportion of be With Knapp-Harts			plained	Adj	R-squared	= -1.54%
ln_cmrdrug	exp(b)	Std. Err.	t	P> t	[95% Conf	. Interval
In cocaine 1	1.382652	.5461887	0.82	0.422	. 6065223	3.15194
cons	1.220997	.2533423	0.96	0.347	.7920386	1.88227

Figure L10. Meta-regression analysis output examining the impact of GBD region (ref. High-Income North America; GBD 1 = Western Europe, GBD 2 = Tropical Latin America) on CMR per 100PY.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-st	tudy variance	e		tau2	= .6'	754
<pre>% residual var:</pre>	iation due to	heterogene:	ity		I-squared res	= 98.9	95%
Proportion of 1	between-study	y variance e	xplained		Adj R-squared	= 1.2	20%
Joint test for	all covaria	tes			Model F(2,13)	= 0	. 97
With Knapp-Har	tung modifica	ation			Prob > F	= 0.40	048
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf.	Interva	al]
In GBD 1	1.689008	.7549208	1.17	0.262	. 6430974	4.435	951
TH CRD 0	2.447457	2.199044	1.00	0.337	.3513243	17.04	499
In GBD 2							

Figure L11. Bubble plot displaying results of meta-regression analysis examining the impact of final year of study follow-up on CMR per 100PY.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-st	tudy varianc	e		tau2	=	. 5979
% residual var:	iation due to	o heterogene	ity		I-squared res	=	98.90%
Proportion of N With Knapp-Hart		•	xplained		Adj R-squared	-	12.54%
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	t <mark>erval]</mark>
finalyear	. <mark>93854</mark> 93	.0341406	-1.74	0.103	.8681087	1	.014706

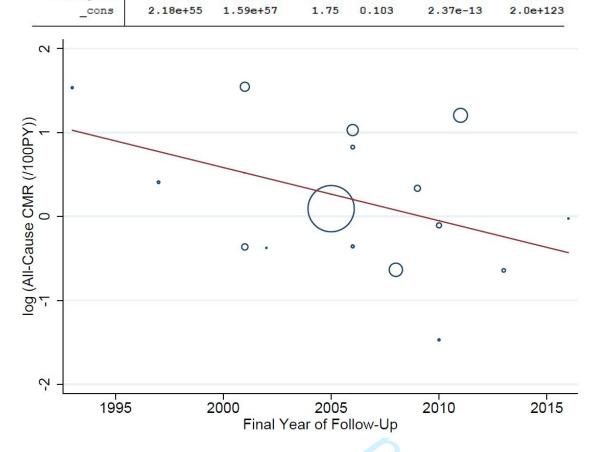


Figure L12. Bubble plot displaying results of meta-regression analysis examining the impact of cohort size on CMR per 100PY.

Meta-regression	n				Number of obs	=	15
REML estimate (of between-st	tudy varianc	e		tau2	=	. 6731
<pre>% residual var:</pre>	iation due to	heterogene.	ity		I-squared res	=	98.70%
Proportion of B With Knapp-Hart			xplained		Adj R-squared	=	-7.82%
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	t <mark>erval</mark>]
ln_cmr sample	exp(b)	Std. Err.	t -0.26	P> t 0.797	[95% Conf.		terval] .000033

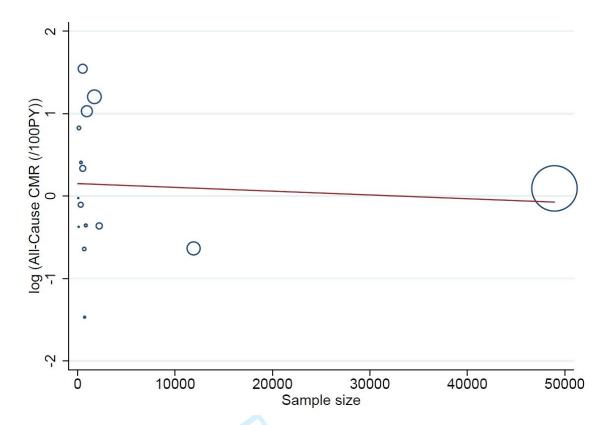


Figure L13. Bubble plot displaying results of meta-regression analysis examining the impact of person-years of follow-up on CMR per 100PY.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-st	tudy variance			tau2	=	.7278
<pre>% residual var:</pre>	iation due to	o heterogenei	ty		I-squared res	=	98.79%
Proportion of 1	between-study	y variance ex	plained		Adj R-squared	=	-6.45%
With Knapp-Hart	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
lengthFU	1.014049	.0376479	0.38	0.713	.9364332	1	.098097
_cons	1.023937	.5541067	0.04	0.966	.3207789	3	.268441

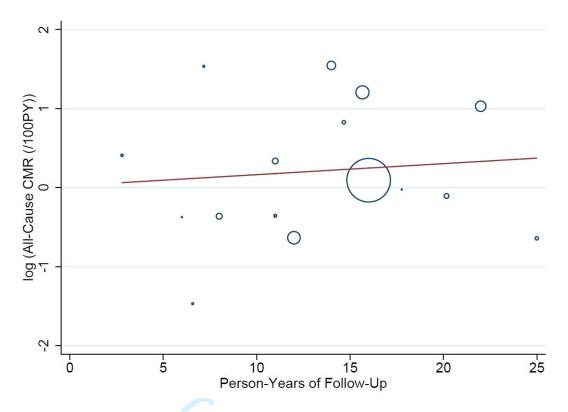


Figure L14. Meta-regression analysis examining the impact of recruitment setting (ref. Treatment clinics and other health services; Setting 1 = Hospital; Setting 2 = Convenience sampling) on CMR per 100PY.

Meta-regression				Nu	mber of obs	=	16
REML estimate of	f between-st	udy variance		ta	12	= .'	7836
<pre>% residual varia</pre>	ation due to	heterogenei	ty	I-:	squared res	= 98	.33%
Proportion of be	etween-study	variance ex	plained	Ad	j R-squared	= -14	. 62%
Joint test for a	all covariat	es		Moo	del F(2,13)	= 1	0.10
With Knapp-Hart	ung modifica	tion		Pro	ob > F	= 0.	9036
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf	. Inte	rval]
_In_setting_1	.7503937	.4920517	-0.44	0.669	.1819971	3.0	93954
In setting 2	.9611738	.4889236	-0.08	0.939	.320292	2.8	84415
	1.319066	.4887616	0.75	0.468	.592 <mark>4</mark> 015	2.9	37088
12							

Figure L15. Meta-regression analysis output examining the impact of study sampling frame (ref. National/Subnational, Area 2 = City) on CMR per 100PY.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-st	tudy variance	e		tau2	=	.4449
<pre>% residual var:</pre>	iation due to	heterogene:	ity		I-squared_res	=	96.64%
Proportion of 1 With Knapp-Har			xplained		Adj R-squared	=	34.92%
ln_cmr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
_In_area_2	2.776357	.9854761	2.88	0.012	1.296724	5	.944331
_cons	.7759408	.1873406	-1.05	0.311	.4623147	1	.302325

Figure L16. Bubble plot displaying results of meta-regression analysis examining the impact of proportion of women within cohort on SMR.

Meta-regression	n				Number of obs	=	9
REML estimate o	of between-st	tudy variance	e <mark>-</mark>		tau2	=	.1861
% residual vari	iation due to	o heterogene:	ity		I-squared_res	=	95.11%
Proportion of H With Knapp-Hart		The state of the s	xplained		Adj R-squared	=	39.96%
ln_smr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
prop_women	.0375019	.0505811	-2.43	0.045	.0015451		9102363
_cons	12.45751	4.452131	7.06	0.000	5.350775	2	9.00318

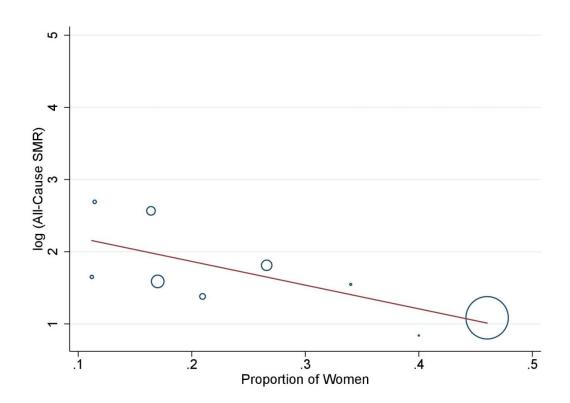


Figure L17. Bubble plot displaying results of meta-regression analysis examining the impact of proportion of people who inject drugs within cohort on SMR.

Meta-regression	1				Number of obs	=	8
REML estimate	of between-s	tudy variance			tau2	=	. 68
<pre>% residual var:</pre>	iation due to	o heterogenei	ty		I-squared_res	=	92.70%
Proportion of B	between-stud	y variance ex	plained		Adj R-squared	=	17.35%
With Knapp-Hart	exp(b)	ation Std. Err.	t	P> t	[95% Conf.	In	tervall
	3.090519	2.472367	1.41	0.208	11		1.88563
1000 - 1000 - 1000 - 100	4.279761	2.78444	2.23	0.067		100	1.02839
_ ^{cons}	4.2/9/01	2.70444	2.25	0.067	.0/10296	2	1.02039

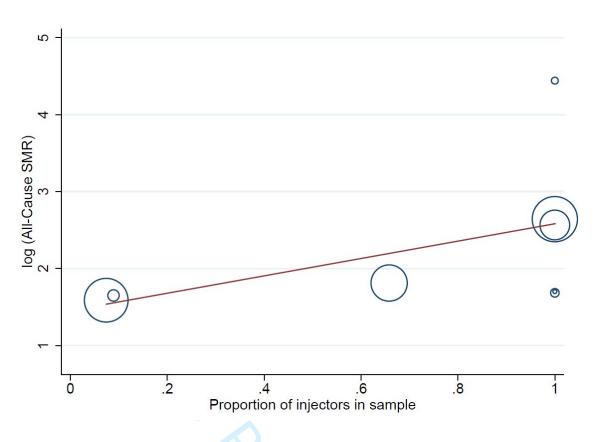


Figure L18. Meta-regression analysis examining the impact of cocaine form used (ref. Cocaine/Cocaine and Crack Cocaine; Cocaine 1 = Cocaine and Heroin) on SMR.

Meta-regression				Nun	ber of obs	=	18
REML estimate or	f between-st	udy variance		tau	12	=	.7204
<pre>% residual varia</pre>	ation due to	heterogeneit	y	I-s	quared_res	=	98.91%
Proportion of be With Knapp-Hart			lained	Adj	R-squared	=	2.50%
ln_smrdrug	exp(b)	Std. Err.	t	P> t	[95% Conf.	. Iı	nterval]
_In_cocaine_1	2.113077	1.400303	1.13	0.276	.5185684	1	8.610423
_cons	6.164663	1.374622	8.16	0.000	3.842519	3	9.890144

Figure L19. Meta-regression analysis output examining the impact of GBD region (ref. High-Income North America; GBD 1 = Western Europe, GBD 2 = Tropical Latin America) on CMR per 100PY.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-s	tudy varianc	e		tau2	=	.8346
<pre>% residual var</pre>	iation due to	o heterogene	ity		I-squared_res	=	99.02%
Proportion of 2	between-stud	y variance e	xplained		Adj R-squared	=	-7.77%
Joint test for	all covaria	tes			Model F(2,13)	=	0.56
With Knapp-Har	tung modifica	ation			Prob > F	=	0.5870
ln_smr	Coef.	Std. Err.	t	P> t	[95% Conf.	In	terval]
_In_GBD_1	207001	.4939925	-0.42	0.682	-1.274207	82	8602048
In GBD 2	.8438148	.9954567	0.85	0.412	-1.306739	2	.994368
_III_GDD_2	.0450140	12201001					

Figure L20. Bubble plot displaying results of a meta-regression analysis examining impact of final year of study follow-up on SMR.

Meta-regression	n				Number of obs	=	16
REML estimate	of between-st	tudy variance	e		tau2	=	. 6437
% residual var.	iation due to	heterogene	ity		I-squared_res	=	98.86%
Proportion of 1	between-study	y variance e	xplained		Adj R-squared	=	16.88%
With Knapp-Har	tung modifica	ation					
ln_smr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
ln_smr finalyear	exp(b)	Std. Err.	t -2.08	P> t	[95% Conf.		terval] 1.00241

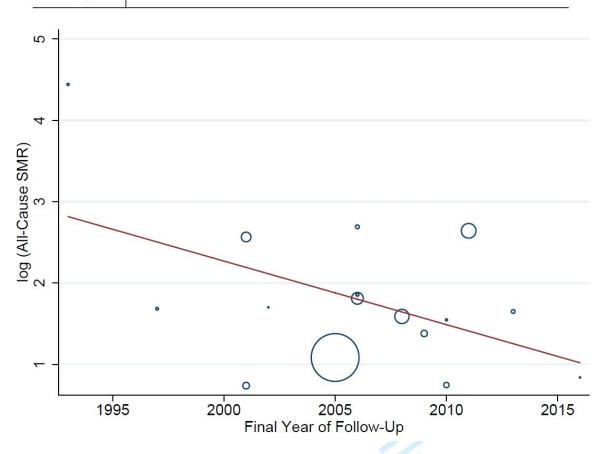


Figure L21. Bubble plot displaying results of a meta-regression analysis examining impact of cohort size of followup on SMR.

Meta-regression	n				Number of obs	=	15
REML estimate of between-study variance % residual variation due to heterogeneity					tau2	=	.3926
					I-squared_res		97.42%
Proportion of h	between-study	y variance ex	xplained		Adj R-squared	=	0.29%
With Knapp-Hart	tung modifica	ation					
ln_smr	exp(b)	Std. Err.	t	P> t	[95% Conf.	In	terval]
ln_smr sample			t -0.98	P> t 0.345	[95% Conf.	00000	terval]

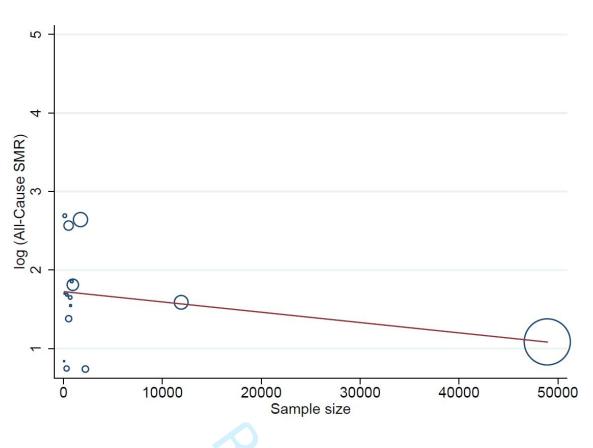


Figure L22. Bubble plot displaying results of a meta-regression analysis examining impact of person-years of follow-up on SMR.

Meta-regression	n				Number of obs	=	16
REML estimate of between-study variance % residual variation due to heterogeneity					tau2	=	.8104
					I-squared_res	=	99.04%
Proportion of 1	between-study	y variance e	xplained		Adj R-squared	=	-4.64%
With Knapp-Hart	tung modifica	ation					
ln_smr	exp(b)	Std. Err.	t	P> t	[95% Conf.	Int	erval]
lengthFU	.9733267	.0381837	-0.69	0.502	.8947816	1.	058767
_ ^{cons}	8.785405	5.019797	3.80	0.002	2.579526	29	92152 .
1							

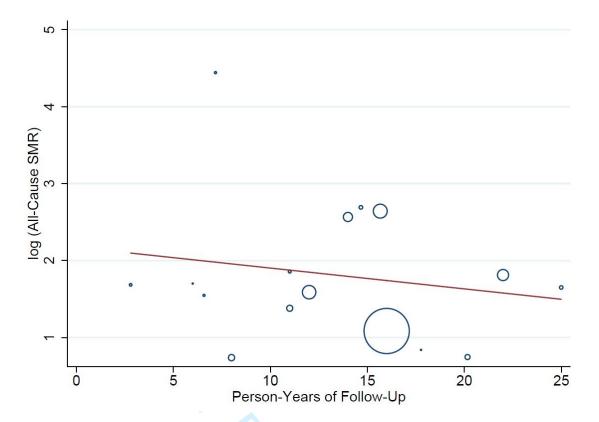


Figure L23. Meta-regression analysis examining the impact of recruitment setting (ref. Treatment clinics and other health services; Setting 1 = Hospital; Setting 2 = Convenience sampling) on SMR.

Meta-regression				Nur	ber of obs	=	16
REML estimate of	tau	=	. 6217				
<pre>% residual varia</pre>	I-s	=	96.99%				
Proportion of be	Adj	=	19.73%				
Joint test for	Model F(2,13) Prob > F			2.83			
With Knapp-Hart				0.0953			
ln_smr	exp(b)	Std. Err.	t	P> t	[95% Conf	. 1	nterval]
In_setting_1	.2472893	.1457852	-2.37	0.034	.0691965		.8837443
_In_setting_1 _In_setting_2	.2472893 .5868937	.1457852	-2.37 -1.17	0.034 0.264	.0691965 .2189067		.8837443 1.573475

Figure L24. Meta-regression analysis output examining the impact of study sampling frame (ref. National/Subnational, Area 2 = City) on SMR.

Meta-regression	n				Number of obs	=	16
REML estimate of between-study variance					tau2	=	.5835
<pre>% residual vari</pre>	I-squared_res	96.65%					
Proportion of k	between-study	y variance e	xplained		Adj R-squared	-	24.65%
With Knapp-Hart	exp(b)	Std. Err.	t	P> t	[95% Conf.	Int	terval]
In_area_2	2.578813	1.042046	2.34	0.034	1.084008		6.1349
cons	3.971081	1.086776	5.04	0.000	2.20796	7	.142107

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