



# Water resources and tourism development in South Asia: an application of dynamic common correlated effect (DCCE) model

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Received: 25 July 2019 / Accepted: 9 March 2020  
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## Abstract

The current study explores the relationship between water resources and tourism in South Asia for the period of 1995–2017. The study employs the CIPS unit root test for stationarity of the variables and the CD test for cross-sectional dependence among cross-sectional units. As for the long-run parameters, a novel technique, known as dynamic common correlated effect (DCCE) model, is used which was recently developed by Chudik and Pesaran (J Econ 188:393–420, 2015b). The outcomes from the DCCE method suggest that water resources have a positive impact on tourism in South Asia. It is also proven that ignoring cross-sectional dependence among the cross-sectional units may bring about misleading outcomes. The findings of the study can be helpful for policymakers to understand the role of water resources in boosting tourism and contributing to the economic prosperity of South Asian countries.

**Keywords** Cross-sectional dependence · PMG · CIPS unit root test · CD test

**JEL classification** L83 · B23 · Q25 · H5

## Highlights

1. The current study explored the relationship between water availability, governance, infrastructure, exchange rate, and tourism in the south Asian context.
2. The study has employed CIPS unit root test and CD test for cross-sectional dependence, while for long-run parameters, a novel technique known as Dynamic common correlated effect (DCCE) model, is used which is recently developed by Chudik and Pesaran (2015b).
3. The DCCE model accommodates heterogeneity, cross-sectional dependence, and stationarity, which is not possible in other long period based econometrical models.
4. It is confirmed from the results that there is a negative relationship between exchange rate and tourism. However, governance, infrastructure, and water availability have a positive relationship with tourism in the South Asian context.
5. It also proves that ignoring cross-sectional dependence among the cross-sectional units may cause misleading outcomes.

Responsible editor: Eyup Dogan

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## Introduction

The tourism industry is one of the important industries in the world. The significance of this industry could be demonstrated by the fact that it increases income, reduces inflation, produces jobs, and develops infrastructure (Jalil et al. 2013; Arain et al. 2019). The tourism industry produced about 10.4% of the world's gross domestic product (GDP) and generated almost 319 million jobs globally or 10% of total world employment in 2018 (Travel & Tourism Economic Impact 2019)<sup>1</sup>. Furthermore, tourism receipts can be considered an export because it implies a source of revenue and improving the balance of payment for a destination country (Meo et al. 2018).

In other words, tourism receipts support the economic development of a country through their positive impact on the economy as a whole. According to the report of the World Tourism Organization by the United Nations, the tourism industry will contribute substantially to the world's GDP by the end of 2050 (UNWTO 2015). Therefore, the development of the tourism industry will be a popular approach for economic growth worldwide (Antonakakis et al. 2015; Rasheed et al. 2019) because a positive growth in the tourism industry will

<sup>1</sup> <https://www.wtcc.org/-/media/files/reports/economic-impact-research/regions-2019/world2019.pdf>

lead to an overall economic development and growth (Lee and Chang 2008). As a result, academicians are now focusing more on the tourism industry.

Existing studies on tourism demand have looked into various determinants of tourism growth, such as GDP, terrorism, inflation, peace index, foreign direct investment (FDI), unemployment, trade openness, climate change, law and order conditions, geopolitical issues, macroeconomic policies, trade, visa policies and bilateral relations with other countries, exchange rate, infrastructure, governance, and water resources (Jalil et al. 2013; Meo et al. 2018; Pablo-Romero and Molina 2013; Fareed et al. 2018; Ali et al. 2020). However, this study employs water resources, exchange rate, infrastructure, and governance, as determinants of tourism in South Asia.

Over the past few decades, it has become evident that the scarcity of water is becoming a threat to society due to a gradual increase in water demand. It must be noted that both tourism demand and the sustainability of any tourist destination depend on the availability of water (Rico-Amoros et al. 2009). Many studies have evaluated water consumption by tourists and concluded that tourists in developed countries consume two to three times more water compared with local water consumption (Holden 2003; Garcia and Servera 2003; Tortella and Tirado 2011). Tortella and Tirado (2011) contended that the water demand of tourists could be a threat to sustainability, generally, in those areas where water is already scarce, and especially in those destinations that are visited by a high percentage of world tourists. Gössling et al. (2012) carried out a quantitative and qualitative study to examine the role of fresh water in the tourism industry. They found that the use of fresh water directly in tourism is less than 1% of world consumption. They also found that indirect requirement for water is greater than direct requirement because of production of food, building materials, and energy. In Bali, Cole (2012) found that fresh water supply/availability is one of the major factors that can boost the tourism industry. Hadjidakou (2014) found that tourism significantly affects economic growth, water resources, and tourism demand. It is therefore evident that water scarcity could adversely affect tourism demand because the availability of water is a necessary determinant for tourism growth. In fact, water resources are a major tourist attraction in any country.

We also employed exchange rate as the most important determinant of tourism growth, following Balaguer and Cantavella-Jorda (2002) and Oh (2005), who argued that the exchange rate variable must be placed along with other explanatory variables in the tourism demand model to avoid the omitted variable issue. There are various studies which found that tourist highly consider exchange rate (Crouch 1995; Patsouratis et al. 2005; Önder et al. 2009). Chadeand and Mieczkowski (1987) analyzed the impact of exchange rate (Canadian-USA) on tourism in Canada, and found that for every 1 % depreciation of the Canadian-US exchange rate,

there is a corresponding 1.26% increase in the number of tourists in Canada. Meo et al. (2018) explored impact of exchange rate (LCU per US\$) on tourism. They found that increase in exchange rate increases tourism demand, while Önder et al. (2009) and Gan (2015) found negative effect of exchange rate (LCU for unit of foreign currency) on tourism. Furthermore, Quadri and Zheng (2010) carried out a study on exchange rate (LCU per Euro) and tourism, and found that in Italy, exchange rate does not affect tourism. Chi (2015) examined impact of exchange rate (US dollar versus foreign currencies) on tourism in USA, and found that appreciation of US dollar negatively affects tourism. It is also observed that effect of exchange rate varies in different tourist destinations.

Currently, tourists are also concerned with the infrastructure and governance of the destination country, both of which have shown significant contribution to tourism inflows (Chingarande 2014). Gunn (1988) and Inskip (1991) concluded that the infrastructure of a tourism destination is a potential attraction for tourists because good infrastructure avails accessibility to different areas of the destination country to tourists. Several researchers have studied tourism products and have concluded that infrastructure is one of the most crucial tourism products that affects tourism demand (Khadaroo and Seetanah 2007, 2008; Seetanah et al. 2011). Getahun and Ayele (2016) found in Ethiopia that infrastructure (water and transportation) affects the tourism industry positively. Mustafa (2019) examined the impact of infrastructure on tourism development in Sri Lanka and found that the impact of infrastructure on tourism was positive in both short-run and long-run. Similarly, in China, Yu (2016) also found a positive relationship between improved infrastructure and tourism industry. The governance of a tourism destination (such as accountability, control of corruption, the rule of law, and regulatory quality) is particularly significant in the selection of a tourism destination and also crucial for sustained growth of the tourism industry (Baggio et al. 2010; Tang & Tan 2015). The tourism sector is a very fragile and vulnerable industry in any country. The way a country is governed has an effect on tourists' willingness to visit that country (Bassil 2014). According to Yap and Saha (2013), tourists are more concerned with their safety and security. Meo et al. (2018) investigated the relationship between the governance index and tourism. They found that governance positively affects tourism demand in Pakistan. Hence, good governance of a destination can attract tourists, and it could also lead to increased tourism demand. Araña and León (2008) found that tourists visit those destinations where they find good governance compared with those with weak governance even if the country has attractive sites. Kim et al. (2018) carried out a panel study of 108 counties to investigate the impact of governance (institutional quality) on tourist inflows. They found that governance positively impacts tourist inflows in developed and emerging economies, and has a greater effect

on tourism in highly developed economies compared with emerging economies. Ghalia et al. (2019) applied the gravity model and found that governance (institutional quality) affects tourism positively. A recent study accomplished by Lee et al. (2020) found that governance (institutional quality) plays a major role in determining tourism in Malaysia.

### Face of tourism in South Asia

South Asian countries (Pakistan, Maldives, Bangladesh, Nepal, India, Bhutan, Sri Lanka, and Afghanistan), with their natural beauty and fascinating tourist sites, attract people from the whole world (Rasul & Manandhar 2009; Ali and Nazar 2017). South Asia is a diverse region, spanning coastal areas to mountains and grasslands to forests, with picturesque beauty, rivers, and diversified climate conditions. For example, Pakistan has Neelum Valley, Shangrila Resort, Naltar Valley, Skardu, Narnan Valley, and Jhelum Valley for tourists. Likewise, Maldives has many tourist attractions, such as Hulhumale Island, Feydhoo, and Veligandu Island. Bangladesh has various historical sites for tourists, for instance, Shahid Minar, Ahsan Manzil, and various beaches. Similarly, Nepal, India, Bhutan, Sri Lanka, and Afghanistan have numerous historical and beautiful places for tourists, but the region's economic conditions are still poor. Despite that, South Asian countries have great potential for tourism, but their share is only 1% of global tourism. The UNWTO (2015) reported that worldwide, tourist arrival was 1133 million but South Asia received just 17.1 million tourists. Considering the potential of tourism in South Asia, this study attempts to analyze the major factors of tourism demand in this region.

The present study contributes to the body of knowledge from various perspectives. Firstly, previous studies on tourism, water resources, exchange rate, infrastructure, and governance have assumed homogeneity across the countries. However, practically, this assumption is not viable (Sadorsky 2013). Secondly, the present study includes water resources as an additional determinant of tourism in the South Asian context; water scarcity is currently an issue in South Asian economies, and earlier studies have ignored this variable. Thirdly, previous studies have ignored the issue of cross-sectional dependence among different countries. If the series suffer from cross-sectional dependence, traditional approaches, such as the GMM model, random-effect model, fixed-effect model, or pooled regression, provide invalid outcomes (Sadorsky 2013; Turkey 2017). Besides, De Hoyos and Sarafidis (2006) found that in the current globalization era, cross-sectional dependence among the economies is not an expectation but a rule to financial integration among the economies. Fourth, the empirical results of this study have significant policy implications for the policymakers who devise policies for boosting tourism in South Asia. Various studies have employed mean group (MG) and pooled mean group (PMG) for the long-term panel data.

However, these methods only consider the heterogeneity problem (Pesaran and Smith 1995) and give misleading results due to the presence of cross-sectional dependence (Al Mamun et al. 2015). Therefore, due to the importance of cross-sectional dependence and heterogeneity, we employed the dynamic common correlated effect (DCCE) model recently introduced by Chudik and Pesaran (2015b). The DCCE model considers both cross-sectional dependence and heterogeneity, and provides results that are more accurate compared with PMG, MG, and AMG models. Hence, the present study is very useful for tourism policymakers in South Asian and other developing economies.

The rest of the paper is organized as follows. The second section elaborates on data and methodology, section three consists of results, and section four gives the conclusion and policy implications.

## Data and methodology

### Variables and data sources

The present study examines the impact of water resources and other control variables (exchange rate, infrastructure, governance) on tourism in South Asian countries, such as India, Pakistan, Bangladesh, Sri Lanka, and Nepal. Due to the unavailability of data, the other South Asian countries are not included. We measured tourism by the total number of tourist arrivals in each of those countries. Numerous empirical studies have used various proxies of water resources, such as Raskin et al. (1997), Falkenmark (1997), Vörösmarty et al. (2000), and Oki and Kanae (2006) used ratio of water consumption to availability; Seckler (1998) used supply of water; Aldaya (2012) used signified ratio of water unavailability with the footprint of water-to-water accessibility; while Falkenmark et al. (2009), Ohlsson and Appelgren (1998), and Falkenmark et al. (1989) employed per capita water resources. Therefore, we also used per capita water availability as a proxy for water resources, while real exchange rate was measured by the change in annual percentage in the official currency rate per USD, following Bano et al. (2018). The governance index<sup>2</sup> was constructed with various governance indicators, and the infrastructure index<sup>3</sup> was calculated by following Palei (2015). Table 1 shows the description of variables and data sources.

<sup>2</sup> We employed principal component analysis (PCA) using reviews to construct governance index of various components of governance from World Bank data (including rule of law, government effectiveness, voice and accountability, absence of violence, political stability, regulatory quality, and control of corruption).

<sup>3</sup> The study constructed infrastructure index following Palei (2015). In this study, infrastructure index was developed by using basic physical infrastructure, i.e., transport infrastructure, telecommunication infrastructure, water sanitation, and energy infrastructure.

**Table 1** Variables description and data sources

Variables	Description	Unit of measurement	Data source
TOR	Tourism development	Number of tourist's arrival	WDI
WR	Water resources	Per capita water availability	FAO
EXCH	Real exchange rate	Based on US dollar exchange rate	WDI
GOV	Governance index	Index	ICRG
INFRA	Infrastructure index	Index	(Palei 2015)

The study used annual data from 1995 to 2017

### Unit root tests in the presence of cross-sectional dependence

A growing body of literature indicated that panel data models are suffered by the issue of cross-sectional dependence in error terms which may arise due to the presence of unobserved components and common shocks that conclusively become the part of spatial dependence, error term, and idiosyncratic pairwise dependence (Pesaran 2004). Moreover, the interaction between economies and financial integration also gives rise to cross-sectional dependence (De Hoyos and Sarafidis 2006). This phenomenon becomes very severe in dynamic panel estimation. In most prior studies, first-generation unit root tests have been used (Maddala and Wu 1999; Levin et al. 2002; Im et al. 2003). These tests, however, have not considered cross-sectional dependence.

On the other hand, the second-generation unit root tests consider the cross-sectional dependencies among panel data (Moon and Perron 2004; Smith et al. 2004; Bai and Ng 2004; Pesaran 2007). In recent studies, third-generation unit root tests have also been reported, which can tackle structural breaks in data. These unit root tests have some drawbacks related to statistical size and power property, as indicated by Hossain (2011). Specifically, the first-generation unit root tests show issues with size properties of data and a false rejection of the null hypothesis in the presence of cross-sectional dependence (Banerjee et al. 2001). Choi (2006) and Pesaran (2007) introduced second-generation unit root tests, which have corrected the issue of false rejection of the null hypothesis. However, the assumption of homogeneity across the cross-sectional dependence is still not true as stated by Urbain and Westerlund (2006).

The present study employed Pesaran's (2004) unit root test since it addresses the issue of false rejection of the null hypothesis in cross-sectional dependence data.

$$a_{it} = \delta_i + \beta_{it}b_{it} + \mu_{it} \tag{1}$$

Equation (1) states the relationship between  $a_{it}$  which depends on residuals  $= \mu_{it}$  and time invariant individual nuisance parameters  $= \delta_i$ . The slopes to be estimated are denoted by  $\beta_{it}$ ,

and  $b_{it}$  refers to the number of regressors. In the subscript, the "i" refers to the cross-section and "t," the time period.

The following hypothesis was tested to investigate the presence of cross-sectional dependence among panel data.

$$H_0 = \rho_{iz} = \rho_{zi} = \text{cor}(\mu_{it}, \mu_{it}) = 0 \text{ for } i \neq z \tag{2}$$

$$H_1 = \rho_{iz} = \rho_{zi} = \text{cor}(\mu_{it}, \mu_{it}) \neq 0 \text{ for some } i \neq z \tag{3}$$

The correlation between the two roots validates CSD as stated in Eqs. 2 and 3 above. The null hypothesis ( $H_0$ ) indicates that there is no cross-sectional dependence among cross-sectional units and vice versa for the alternate hypothesis ( $H_1$ ).

### Panel cointegration

A literature review indicates that cointegration techniques have been widely used in empirical studies. These cointegration techniques have historically been under strict scrutiny for long-series data estimation (Pedroni 1997). Many studies have reported that these cointegration techniques exhibit more problems in application concerning period and less concerning data frequency (Shiller and Perron 1985; Perron 1991). In this research, we used the bootstrap cointegration approach introduced by Westerlund and Edgerton (2007) to explore the long-term relationship among the variables under study. This approach is more efficient than other cointegration techniques. For example, contrary to Pedroni's (2004) cointegration approach, it takes structural breaks into account as well. Moreover, this method considers lead-lag lengths when applied to short-duration data (Persyn and Westerlund 2008).

Equation (4) refers to the bootstrap panel cointegration approach proposed by Westerlund and Edgerton (2007).

$$\Delta y_{it} = \delta' d_t + \alpha_i \left( y_{i,t-1} - \beta_i' x_{i,t-1} \right) + \sum_{j=-qi}^{qi} a_{ij} \Delta y_{i,t-1} + \sum_{j=-qi}^{qi} \gamma_{ij} \Delta x_{i,t-1} + e_{i,t} \tag{4}$$

The above equation shows the relationship for the endogenous variable  $= \Delta y_{it}$ , for three various cases,  $d_t$  refers to the deterministic component. Subscript  $t$  and  $i$  denote time-period and cross-sectional units, respectively.



### Dynamic common correlated effects

The literature review shows that researchers have not considered cross-sectional effects and have only worked with homogenous slopes in earlier studies. The related studies show various techniques for panel data analysis, including GMM, random-effect, and fixed-effect models. In these models, only the intercept changes among the cross-sectional units, leaving a high degree of homogeneity. This assumption is not true and produces misleading results (Turkay 2017).

Due to these reasons, panel data estimation with heterogeneous coefficients among cross-sectional units over longer periods has attracted the attention of researchers in the recent past (Pesaran and Smith 1995; Pesaran et al. 1999). Researchers, the world over, have also paid attention to the issue of the dependence of cross-sectional units (Pesaran 2006; Chudik and Pesaran 2015b). In this work, we applied the dynamic common correlated effect (DCCE) approach introduced by Chudik and Pesaran (2015b). This approach uses the principles of PMG estimation by Shin et al. (1999), MG estimation of Pesaran and Smith (1995), CCE estimation of Pesaran (2006), and the estimation by Chudik and Pesaran (2015a). Blackburne and Frank (2007) suggested xtpmg command (pooled mean group estimator) for non-stationary and heterogonous large data sets, as the PMG estimator does not consider the cross-sectional dependence. Eberhardt (2012) used common correlated effects without pooled coefficients or DCCE. The CCE estimation also does not consider the lag value of an endogenous variable as an explanatory variable (Chudik & Pesaran 2015), while the DCCE method considers homogenous and heterogeneous coefficients and the DCCE and also takes into account cross-sectional dependence. This technique incorporates heterogeneous slopes and cross-sectional dependence by taking cross-sectional means and lags into consideration.

Moreover, this method works well for small sample size by using the jack-knife correction approach (Chudik & Pesaran 2015). Another major benefit of this technique is its estimation robustness in the presence of structural breaks in data (Kapetanios et al. 2011). This technique also performs satisfactorily for unbalanced panel data (Ditzen 2016).

We used the dynamic equation of the DCCE model from Chudik and Pesaran (2015b) as follows:

$$WR_{it} = \alpha_i WR_{it-1} + \delta_i x_{it} + \sum_{p=0}^{P_T} \gamma_{xip} \bar{X}_{t-p} + \sum_{p=0}^{P_T} \gamma_{yip} \bar{Y}_{t-p} + \mu_{it} \quad (6)$$

For our work, in Equation (6), WR refers to water resources,  $\alpha_i WR_{it-1}$  is the lag of WR as an independent variable,  $\delta_i x_{it}$  refers to the set of independent variables, and  $P_T$  is the limit of lags included in the cross-section averages.

**Table 2** Descriptive statistics

	TOR	WR	EXCH	GOV	INFRA
Mean					
Median	1,528,207	3519.44	69.25	-0.07	9.46
Maximum	6,968,000	2549.00	129.06	-0.52	10.01
Minimum	125,000.0	1391.00	36.31	-2.63	5.02
Std. dev.	1,976,441	2802.96	25.95	0.27	3.06
Skewness	1.709536	1.14	0.92	0.27	0.03
Kutosis	4.657334	2.48	2.86	1.87	2.50
Jarque-Bera	17.44447	6.61	4.14	1.89	0.30
Probability	0.000163	0.03	0.12	0.38	0.85

### Results and discussion

In recent empirical research, cross-sectional dependence and macroeconomic determinants have received greater attention. Currently, cross-sectional dependence is considered a basic rule and not an expectation between the cross-sectional units (Turkay 2017). To prevent cross-sectional dependencies between the cross-sectional units and misleading parameters (Hsiao and Tahmiscioglu 2008), we employed the Pesaran (2004) test of cross-sectional dependence (CD) by considering the cross-sectional dependence between the cross-sectional units' significances. Furthermore, the test was done on pairwise mean correlated ordinary least square (OLS) residuals of the individual regression in a panel. The results of descriptive statistics are shown in Table 2 and the results of cross-sectional dependence are provided in Table 3. The CD test's null hypothesis represents the independence in cross-sectional units or non-cross-sectional dependence. The findings in Table 3 confirm the existence of cross-sectional dependence between the units of the cross-section.

Due to the importance of the cross-sectional dependence among the variables in the globalization era, the CD test was employed in the study and the results are presented in Table 3. It is observed that the data series experience cross-sectional dependence. Another significance of the CD test is that it is helpful for knowing whether the first-generation panel unit root tests (Levin et al. 2002; Im et al. 2003) or second-

**Table 3** Cross-sectional dependence test results

	CD Test	p value
TOR	15.87	0.00*
WR	9.42	0.00*
EXCH	7.87	0.00*
GOV	7.41	0.00*
INFRA	8.52	0.00*

\*Refers to the rejection of null hypothesis of CD at 1% level of significance

**Table 4** First-generation unit root tests (LLC & IPS)

	Levin, Lin, and Chu				Im, Pesaran, and Shin W-stat			
	Level		1st diff.		Level		1st diff.	
	Stat.	Prob.	Stat.	Prob.	Stat.	Prob.	Stat.	Prob.
TOR	-1.12	0.41	-12.74	0.00*	-1.08	0.13	-11.41	0.00*
WR	0.32	0.62	-2.90	0.00*	-0.23	0.40	-2.85	0.00*
EXCH	-1.24	0.10	-4.85	0.00*	-2.57	0.00*	-6.19	0.00*
GOV	-2.37	0.00*	-7.07	0.00*	-0.72	0.23	-8.76	0.00*
INFRA	-5.43	0.00*	-6.50	0.00*	-4.59	0.00*	-7.17	0.00*

\*Refers to level of significance at 1%

generation unit root techniques (Pesaran 2007; Chang 2004) are appropriate for the proposed variables of the study. Cross-sectional dependence is accommodated by first-generation unit root tests, and they assume homogeneity across the cross-sections. However, second-generation unit root tests consider cross-sectional dependence (Kahia et al. 2016). Hence, this study employed both first and second-generation unit root tests to prevent misleading inferences. The findings of first-generation unit root tests are shown in Table 4, and the findings of the second-generation unit root tests in Table 5. All the variables are stationary at the first difference and level as shown in Table 4.

Furthermore, this study employed the second-generation Pesaran’s (2007) CIPS unit root test that allows cross-sectional dependence among the series and provides more accurate results compared with first-generation unit root tests. The findings of CIPS unit root test are presented in Table 5 that confirm that the governance index is stationary at level, while tourism, exchange rate, infrastructure, and water resources are stationary at first difference. None of the variables is stationary at the second difference according to both first and second-generation unit root tests. Hence, we proceeded with the DCCE method to estimate the long-run association among the proposed variables.

In Table 6, the findings of Pedroni’s cointegration tests are shown which confirm the absence of a long-run association among the purposed variables. Westerlund and Edgerton (2008) found that most cointegration tests cannot account for structural breaks as these tests may provide misleading results. However, Westerlund and Edgerton’s (2008) test of cointegration deals with cross-sectional dependence, structural breaks, serial correlation, and heteroskedasticity issues and provides robust outcomes compared with traditional cointegration tests. In Table 7, the Westerlund (2007) bootstrap panel cointegration results are presented which confirm the presence of long-run association or cointegration between TOR, WR, EXCH, GOV, and INFRA. The probability values

<sup>4</sup> We employed xtwest command for Westerlund cointegration.

**Table 5** Second-generation unit root test (CIPS)

	Level	First difference
TOR	-1.96	-6.17*
WR	0.29	-3.00*
EXCH	-1.91	-5.66*
GOV	-2.16***	-3.82**
INFRA	-1.21	-5.68*

The symbols \*, \*\*, and \*\*\* refer to level of significance at 1%, 5%, and 10% respectively

of Gt, Ga, Pt, and Pa (Persyn & Westerlund (2008)<sup>4</sup>) cointegration tests are less than 0.05 that reject the null hypothesis of no cointegration.

The findings of the PMG model in Table 8 confirm that water resources (WR), exchange rate (EXCH), and infrastructure index (INFRA) have a negative and significant relationship with tourism. It means an increase in water resources, exchange rate, and infrastructure index decreases tourism development in South Asia. However, the governance index (GOV) is positively and significantly linked to tourism. It means an increase in governance leads to an increase in tourism in South Asia.

Considering the diagnostic issues of the model, the CD test confirms that series suffer from cross-sectional dependence, while the PMG model does not accommodate the cross-sectional dependence issue. Therefore, the PMG approach could provide misleading outcomes. Hence, considering the issues with PMG model, the study also employed the DCCE model. Table 9 shows the results of the DCCE model. It is found that exchange rate using the DCCE model has a negative and significant relationship with tourism. However, we found a major change in coefficients using the DCCE model. We found that water resources, governance index, and infrastructure index are positively and significantly linked to tourism development. It means an increase in water resources, governance index, and infrastructure index leads to an increase in tourism in South

**Table 6** Pedroni residual cointegration tests

	t-stat	Prob.	Weight t-stat	Prob.
H1: common coefficients (within dimensions)				
V-stat	-0.51	0.69	-0.57	0.71
Rho-stat	0.67	0.74	0.75	0.77
PP-stat	-0.67	0.24	-0.26	0.39
ADF-stat	-0.35	0.36	-0.08	0.46
H1: individual coefficients (between dimensions)				
Rho-stat	1.46	0.92		
ADF-stat	-1.12	0.13		
PP-stat	-1.77	0.03**		

\*\*Refers to level of significance 5%

**Table 7** Westerlund ECM panel cointegration tests

H0: no cointegration	Value	Robust <i>P</i> value
Gt	- 5.567	0.000*
Ga	- 24.168	0.000*
Pt	- 12.363	0.000*
Pa	- 23.983	0.000*

\*Refers to level of significance at 1%

Asia. The findings of the study are aligned with Rico-Amoros et al. (2009), who argued that water resources play a vital role in tourist destination selection. Meo et al. (2018) found that the exchange rate negatively affects tourism demand. They also found that governance and tourism are linked positively and significantly because good governance motivates tourists to visit a destination due to security. Khadaroo and Seetana (2008) also found that infrastructure plays a vital role in attracting tourists. However, the DCCE model provides more accurate results compared with the PMG or other cointegration-based long data set panel approaches.

### Conclusion and policy

The current study explores the relationship between water resources and tourism in the South Asian context using a novel technique, known as the dynamic common correlated effect (DCCE) model. The study used annual data from 1995 to 2017. From a theoretical aspect, this study makes a contribution, in that, it considers cross-sectional dependence among cross-sectional units (using the DCCE approach), which means during this era of globalization, the countries suffer a lot from changes in some other countries. In addition, we incorporated water resources in the tourism model as one of the major tourist attraction factors, which has not been discussed empirically in earlier studies by incorporating cross-sectional dependence. The study confirms a positive and significant effect of water resources on tourism. Therefore, based on the findings, it can be claimed that sufficient water resources can enhance the tourism sector. The tourism sector depends on water resources. Therefore, water scarcity, poor quality of water resources, or media news about

**Table 8** Results of pooled mean group (PMG) estimation

Exog. var	Coefficient	<i>P</i> value
WR	- 0.02	0.00*
EXCH	- 0.07	0.86
GOV	0.44	0.00*
INFRA	- 0.15	0.01*
C	16.09	0.00*

\*Refers to level of significance at 1%

**Table 9** Results of dynamic common correlated effect (DCCE) estimation

Exog. var	Coefficient	<i>p</i> value
TOR (- 1)	- 0.09	0.02**
WR	0.15	0.00*
EXCH	- 0.09	0.02**
GOV	0.01	0.01*
INFRA	0.11	0.00*

The symbols \* and \*\* refer to level of significance at 1% and 5% respectively

the water crisis in South Asia can hugely harm the tourism industry (Hall 2010; Hall and Stoffels 2006). The policymakers should devise policies for the tourism industry by considering water management. Currently, South Asia is facing water issues, and in the near future, water will be a scarce natural resource in South Asia. Furthermore, based on analysis of the UNWTO (2015), the tourism industry will grow in the coming decades, and some regions will face water scarcity issues if reasonable measures are not taken. Therefore, for sustainable tourism, water management requires proper attention. Furthermore, water management is needed not only at country level but also at individual level. Companies, and more importantly, tourism-related industries, such as the hotel industry that use water for their production, must have a solid plan for water-saving. In hotels, there must be proper plans for water usage in bathrooms, for laundry, swimming pools, gardens, the kitchen, and for housekeeping.

The current study suffers from one limitation, i.e., we only targeted South Asia. Due to the unavailability of data, a few countries were excluded. In future, a research can be conducted to compare developing and developed nations. Furthermore, research should incorporate other natural factors, such as effects of temperature and rain in the tourism model using quantile on quantile regression or quantile ARDL model that can provide more accurate results.

**Acknowledgments** We would like to acknowledge the contributions of research participants to this paper.

**Contribution** All the authors contributed equally.

### Compliance with ethical standards

**Conflict of interest** The authors declare that they have no conflict of interest.

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