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Age-related differentiation in verbal and visuo-spatial working memory processing in childhood

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Abstract

Working memory (WM), a key feature of the cognitive system, allows for maintaining and processing information simultaneously and in a controlled manner. WM processing continuously develops across childhood, with significant increases both in verbal and visuo-spatial WM. Verbal and visuospatial WM may show different developmental trajectories, as verbal (but not visuospatial) WM relies on internal verbal rehearsal, which is less developed in younger children. We examined complex VWM and VSWM performance in 125 younger (age 4 to 6 years) and 101 older (age 8 to10 years) children. Latent multi-group modeling showed that (1) older children performed better on both verbal and visuospatial WM, and (3) a model with two factors representing verbal and visuospatial WM fit the data better than a one-factor model. Importantly, the correlation between the two factors was significantly higher in younger than in older children, suggesting an age-related differentiation of verbal and spatial WM processing in middle childhood. Age-related differentiation is an important characteristic of cognitive functioning and thus the findings contribute to our general understanding of WM processing.

Keywords: Working memory; Early childhood; Middle childhood; Age-differences; domain differentiation

Age-related differentiation in verbal and visuo-spatial working memory processing in childhood

Working memory (WM) is a key feature of our cognitive system, allowing for the simultaneous maintenance and processing of information in a controlled manner (Baddeley & Hitch, 1994). WM is an excellent predictor for academic development in childhood, as it helps children to process complex tasks, update and maintain task goals, and remember and follow instructions (Alloway, Gathercole, Kirkwood, & Elliott, 2009). Consistently, WM, which improves most during the preschool period but continues to develop well into adolescence (Gathercole, Pickering, Ambridge, & Wearing, 2004), supports learning across academic domains, including reading and mathematics (Alloway et al., 2005; Alloway, Gathercole, Willis, & Adams, 2004; Cowan, 2014, and Titz & Karbach, 2014, for reviews). Several WM models are currently existing and discussed (e.g., Baddeley, 2000; Barrouillet & Camos, 2012; Oberauer, 2009). They all contributed substantially to our understanding of WM and agree on the basic assumptions that WM capacity is limited and that reliable individual differences in this capacity exist (e.g., Oberauer, 2009). They do, however, differ in the assumptions about the components of WM. For example, there is debate of whether verbal and visuospatial information are processed and stored in different components (e.g., Alloway, Gathercole, & Pickering, 2006). One way to investigate this is by examining the developmental trajectory of verbal and visuospatial WM. Among the most commonly used tasks to assess WM abilities in childhood are complex WM span tasks. In WM span tasks, participants are presented a sequence of stimuli (encoding phase) and are asked to repeat them in the correct or reversed order after a short delay (recall phase). Complex WM span tasks increase demands on executive functioning by adding a processing task to the encoding phase, which requires attention and thus prevents it from contributing to storage during processing episodes. For instance, participants are asked to decide for every stimulus they see whether it is large or small while encoding the sequence. In large cross-sectional datasets, research repeatedly demonstrated substantial age-related WM

performance improvements across childhood on various WM span tasks (e.g., Michalczyk, Malstädt, Worgt, Könen, & Hasselhorn, 2013).

While different components of WM are clearly separable in adults (e.g., Jarrold & Towse, 2006; Vallar & Papagno, 2002), the extent to which these components differentiate during childhood remains to be clarified. Findings related to Baddeley's model demonstrated that the phonological loop and the visual-spatial sketchpad seem to be at least partially separable in both 4- to 8-year-olds (Pickering, Gathercole, & Peaker, 1998; Roebers & Zoelch, 2005) and 11- to 14-year-olds (Jarvis & Gathercole, 2003). However, this differentiation is still not as well-established as the one between the two slave systems and the central executive (Alloway et al., 2006; Gathercole et al., 2004), indicating that the phonological loop and visuospatial sketchpad may continue differentiating across middle childhood. Consistently, performance on tasks requiring the phonological loop or the visuospatial sketchpad correlates with different outcomes, i.e., performance on verbal span tasks significantly correlates with measures of attainment levels in linguistic domains (e.g., English), whereas performance on spatial span tasks significantly correlates with mathematics and science measures (Jarvis & Gathercole, 2003).

Therefore, our first aim was to clarify whether verbal and visuo-spatial working memory processing differentiates during childhood by comparing performance on complex verbal WM (VWM) and visuo-spatial WM (VSWM) span tasks in 4- to 6-year-olds and 8- to 10-year-olds.

Although research has repeatedly demonstrated substantial age-related WM performance improvements across childhood on various WM span tasks (e.g., Michalczyk et al., 2013), it is unclear whether children's performance on complex span tasks is influenced by the verbal or visuospatial nature of the to be recalled information. Young children may perform better on VSWM compared to VWM complex span tasks, because they have had less practice on verbal rehearsal strategies. For instance, Flavell, Beach, and Chinsky (1966) tested 5-, 7-, and 10-yearolds on a serial recall task and results show that the use of rehearsal strategies increased with age and improved performance of item recall. Whereas younger children only passively used rehearsal and mostly just repeated a single item, older children and adults actively used rehearsal strategies, e.g., repeating a series of items (e.g., Cuvo, 1975; Hitch, Halliday, & Littler, 1993; Ornstein, Naus, & Stone, 1977). Further evidence for less practice in rehearsal with age comes from studies with adults, showing that articulatory suppression during item presentation significantly reduced recall performance, leading to recall patterns similar to those found in 5year-olds (Cowan, Cartwright, Winterowd, & Sherk, 1987). Direct evidence from studies with children showed that long and similar item names resulted in decreased rehearsal accuracy in 11year-olds (only based on VWM span tasks; Hitch, Halliday, Schaafstal, & Heffernan, 1991). Once children get older and learn to effectively use rehearsal strategies, performance differences between VWM and VSWM may disappear. However, recent evidence based on directly comparable tasks tapping both VWM and VSWM within the same sample is missing. Thus, the present study addressed this question by directly comparing children's performance on a complex VWM span task and a structurally comparable complex VSWM span task.

In the current study we investigated two questions. (1) Do verbal and visuospatial WM progressively differentiate with age? (2) Do younger children show asymmetrical performance between the verbal and visuospatial domains that no longer exists in older children? To investigate these questions, we examined children in two age groups (younger: 4-6 years and older: 8-10 years) that performed two complex WM span tasks, one verbal and one visuospatial task. If VWM and VSWM progressively differentiate with age, we expect (1) a stronger correlation between two latent factors representing VWM and VSWM in the younger compared to the older children. Moreover, if younger children are less experienced using verbal rehearsal than visuospatial WM processing strategies, we expect (2) that they should perform better on the VSWM span task than on the conceptually comparable VWM span task and that this difference should be reduced or non-significant in older children.

Method

Participants

The sample comprised a total of 226 children, 125 children aged 4 - 6 years ($M_{Age} = 59.7$ months, SE = 0.59; 70 boys) and 101 children aged 8 - 10 years ($M_{Age} = 106.4$ months, SE =

0.64; 59 boys). Eleven additional children were excluded from the analyses because they did not finish both tasks (4), were diagnosed with ADHD, ASD or epilepsy (6) or because of technical errors (1). Half of the data in both age groups were collected in Germany and the other half in Scotland. The children were either tested in the laboratory or in a quiet room in their kindergarten or daycare. Participants were randomly assigned to start with one of the two adaptive complex working-memory span tasks (VWM and VSWM). Informed parental consent was obtained and children received a certificate for participation.¹

Materials, design and procedures

The tasks were presented on laptop computers with the software package Eprime 2.0 (http://www.pstnet.com/products/Eprime/; Schneider, Eschmann, & Zuccolotto, 2002). The test session started with a short game to familiarize the children with the use of the computer mouse. Afterwards, they were introduced to the WM tasks.

Adaptive complex *verbal working-memory (VWM) span task*: In the encoding phase, children heard a series of words corresponding to food items while watching a picture of that item on the screen (e.g., bread, milk...). After each word, children indicated for each picture if it was large or small by clicking on the corresponding response buttons in the lower left and right corners of the screen. At the end of the series, they saw an array containing 8 food items, including those that had just been presented, and were asked to reproduce the sequence of items in the correct order by clicking on the corresponding pictures. After each click, a grey dot was presented in the selected location for 500 ms to indicate that the click had been registered. Children performed four practice trials followed by 12 test trials. To ensure comparability between countries, we used one-syllable words in English as well as in German as stimuli. Further, we considered the age of acquisition of the words (Kuperman, Stadthagen-Gonzalez, & Brysbaert, 2012) to ensure that all children were familiar with the stimuli (e.g., bread, milk, salt, egg, tea, juice, nuts, rice).

Adaptive complex *visuo-spatial working memory (VSWM) span task:* The task procedure was identical to the verbal WM span task, except for the following differences. Children saw an

apple presented at different locations in a 3x3 grid. For each apple, they were asked to indicate if the apple was complete or if bites were taken from it by clicking on the corresponding response button in lower left and right corners of the screen. At the end of the series, children were asked to reproduce the sequence of locations by clicking on the corresponding grid.

In order to control for between-group differences, we also assessed children's verbal abilities and processing speed. Analysis of variance (ANOVA) showed that there neither were differences in terms of language abilities and processing speed between the children tested in Germany and Scotland nor between those performing the VWM or the VSWM task first. However, older children outperformed younger ones on both tasks (see Supplementary Material for details on the tasks, descriptive statistics, reliabilities and ANOVA results).

Task difficulty adjustment and scoring

Task difficulty (i.e., number of items to be remembered) was individually adapted as a function of each child's performance, starting with one item in the first trial and continuing for a total of 12 trials (see Studer-Luethi, Kodzhabashev, Hogrefe, & Perrig, 2015, for a similar procedure). If children responded correctly to both the WM span task and the processing task, task difficulty was increased by one item in the following trial. If they responded incorrectly to both tasks or to the span task only, task difficulty was decreased by one item in the next trial. If they responded correctly to the span task but incorrectly to the processing task, task difficulty did not change. We calculated two scores for each child and each WM span task (VWM and VSWM) representing average (average trial length) and maximum (maximum correct trial length) performance.

Statistical modeling procedures

All models are confirmatory factor models (CFA) based on maximum likelihood estimation with robust standard error correction (MLR; *Mplus* 7.4). For model identification, the first factor loading of each latent variable was fixed to one. Model fit was evaluated with the χ^2 test, the Comparative Fit Index (CFI), the Root Mean Square Error of Approximation (RMSEA), and the Standardized Root Mean Square Residual (SRMR) in reference to Beauducel and

Wittmann (2005). We tested measurement invariance by comparing different levels of invariance (configural, metric, scalar, and strict, e.g. Cheung & Rensvold, 2002). None of the variables had missing cases.

Results

Descriptive statistics and reliabilities are shown in Table S1 and correlations of all measures are shown in Table S2. We tested the measurement invariance of VWM and VSWM performance separately for age group (age 4–6 vs. 8–10), country (Germany vs. Scotland), and sex (female vs. male). In each case, invariance was evaluated in a two-group CFA with two factors (VWM and VSWM) having two indicators each (average and maximum performance; cf. Figure 1 for an example). For all cases, the fit indices of all invariance levels are included in Table 1. Overall, all models fitted the data perfectly (all χ^2 with *ps* >.11) and eight out of nine χ^2 -difference tests were non-significant (indicating that the assumption of a higher level of measurement invariance for age-group, which resulted in a significantly higher χ^2 but not in a severe drop in the CFI (Δ CFI = .003), indicating that there is a small but not severe violation of measurement invariance (e.g. Cheung & Rensvold, 2002). Thus, overall, the data supported the assumption of measurement invariance (e.g. Cheung & Rensvold, 2002). Thus, overall, the data supported the assumption of measurement invariance over age group, country, and sex (Table 1), allowing for comparisons of means and variances between these groups (e.g., Steenkamp & Baumgartner, 1998).

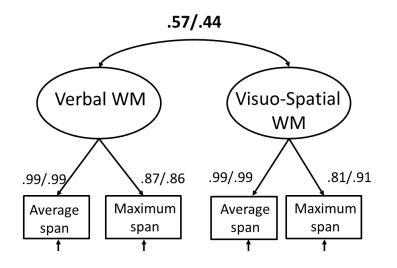


Figure 1. Multi-group model with two age-groups (younger/older children) and strict measurement invariance across age-groups. Read from top to bottom the digits represent latent factor correlations and standardized factor loadings, respectively, for both age groups 4–6/8–10 years.

In a CFA model comparison, we found that two factors for WM-performance (VWM and VSWM) explained the data better than one factor (WM) for both the younger and older children. Two WM-factors fitted the data well in both the younger [$\chi^2(df=2)=0.49$, p=.79; CFI =1; RMSEA = 0 (90 % CI =0-.12); SRMR <.01] and older age-group [$\chi^2(df=2)=0.15$, p=.93; CFI =1; RMSEA = 0 (90 % CI =0-.06); SRMR <.01]. Testing their correlation against 1, however, resulted in a severe drop of model fit in both age-groups [Younger: $\Delta \chi^2(df=1)=83.06$, p<.01; Δ CFI =.28; Older: $\Delta \chi^2(df=1)=157.99$, p<.01; Δ CFI =.41], indicating that two factors represent the data better than one. This is not surprising when considering the strength of the observed correlations between the factors of .57 in the younger and .44 in the older age group. Using a two-group CFA (Figure 1) with strict measurement invariance across age-groups and perfect model fit [Table 1, line 4], we found that both correlations were significantly different from each other [$\Delta \chi^2(df=1)=9.60$, p<.01; Δ CFI =.01], indicating that the relation between VWM and VSWM was higher in younger children than in older children.

Using this two-group CFA (Figure 1) with strict measurement invariance across age groups, we compared latent factor means (see Table S3). Performance on both WM factors was significantly higher for older children with very large effect sizes for both factors. Further, performance on VWM was significantly better than on VSWM for both age groups. This effect was large in the younger and medium-sized in the older age group (see Table S3).

Table 1

Comparison	Invariance Level	r² (đf)	CEI	CFI RMSEA (90% CI) SRMR	SRMR	SB-Scaled $\Lambda \gamma^2$ (d)	A CFI
		10.1 1					
	1. Configural	0.61(6), <i>p</i> = .99	1.00	0 - 0) 0	.004		
Age-Group (Age 4-6 vs.	2. Metric	10.17(8), <i>p</i> = .25	700.	.05 (013)	.112	$^{1.2}$ 8.50(2), $p = .01$.003
Àge 8-10)	3. Scalar	10.11(10), <i>p</i> = .43	1.00	.01 (010)	.109	2,3 0.12(2), $p = .94$	003
	4. Strict	12.08(12), <i>p</i> = .44	1.00	.01 (010)	.102	3,4 1.96(2), $p = .38$	0
Country	1. Configural	5.60(6), <i>p</i> = .47	1.00	0 (012)	.010		
(Germany vs. Scotland)	2. Metric	9.76(8), <i>p</i> = .28	666	.04 (012)	.033	1,2 4.28(2), $p = .12$.001
Ì	3. Scalar	12.13(10), <i>p</i> = .28	866	.04 (012)	.039	2,3 2.37(2), $p = .31$.001
	4. Strict	15.25(12), <i>p</i> = .23	799.	.05 (011)	.046	3,4 3.30(2), $p = .19$.001
	 Configural 	10.21(6), p = .12	966	.08 (016)	.011		
Sex (Female vs.	2. Metric	10.57(8), <i>p</i> = .23	866	.05 (013)	.013	$1.2 \ 0.19(2), p = .91$	002
male)	3. Scalar	10.75(10), <i>p</i> = .38	666	.03 (011)	.014	2,3 0.23(2), $p = .89$	001
	4. Strict	13.95(12), <i>p</i> = .30	866	.04 (011)	.020	^{3,4} 3.46(2), <i>p</i> = .18	.001

Tests of measurement invariance for verbal and visuo-spatial working-memory performance in children

CFI = Comparative Fit Index; RMSEA = Root Mean Square Error of Approximation; CI = Confidence Interval; SRMR = Standardized Root Mean group CFA with two factors (VWM and VSWM) having two indicators each (average and maximum performance; cf. Figure 1 for an example). Notes. We tested measurement invariance separately for age-group, country, and sex. For each case, invariance was evaluated in a two-Square Residual; SB-scaled = Satorra–Bentler scaled; $\Delta \chi^2 = \chi^2$ -Difference test.

Discussion

In the current study, we investigated younger and older children's performance on adaptive complex VSWM and VWM span tasks. Specifically, we tested (1) whether VWM and VSWM progressively differentiate with age, and (2) whether younger children show performance differences between VWM and VSWM span tasks that are no longer evident in older children. Results showed that older children performed significantly better than younger children in both span tasks, replicating former findings (e.g., Gathercole et al., 2004; Michalczyk et al., 2013). Using confirmatory factor modeling, we found that VSWM and VMW were significantly correlated. Critically, the correlation was significantly higher in younger than in older children². Moreover, both age groups performed significantly better in the VWM than in the VSWM span task, which is inconsistent with previous data (e.g., Cowan et al., 1987).

The current findings show that working memory processes in the verbal and visuospatial domains are already partially separable in children, which is consistent with the distinction between the phonological loop and visuospatial sketchpad proposed by Baddeley (Baddeley & Hitch, 1994). Our findings, however, are also consistent with other models of WM that do not include distinct visuo-spatial resources (e.g., Barrouillet & Camos, 2012; Oberauer & Lewandowsky, 2013). Such recent models conceive of a separable verbal component, but take visuospatial processing within the general WM resource. For instance, the latest time-based resource-sharing (TBRS) model of Barrouillet and Camos (2014) differs from Baddeley's model in that it does not include a visuo-spatial WM sub-system. This model includes a verbal rehearsal component and a general resource (coordinating attentional refreshing), but nothing comparable to the verbal rehearsal component in the visuo-spatial domain. Consistently, there is evidence suggesting that visuo-spatial measures may be part of the general WM resource (i.e., sustained attention; Vergauwe, Barrouillet, & Camos, 2010, see also Gray et al., 2017), while verbal WM seems to be a distinct component (e.g., Kane et al., 2004; Park et al., 2002). The separable but correlated VWM and VSWM processes found in the current study are both consistent with the WM structure in Baddeley's model (i.e., representing the phonological loop and the visuo-spatial

sketchpad) and WM structure proposed by Barrouillet and Camos (i.e., representing the verbal rehearsal component and the general resource). However, although the VWM task primarily relied on verbal rehearsal, it included visual stimuli. Thus, one may argue that it was possible to rely on visual processes instead of verbal rehearsal to solve the task. If children would have mainly used such a visual strategy - even though visuospatial processing is more demanding than verbal processing - this would have contributed to a one-factor solution. Given that our data clearly showed a two-factor model, we consider it unlikely that visual strategies were the primary means to solve the task. Although the two tasks used in this study do not allow distinguishing between general and domain-specific variance in the model, higher correlations between the complex VWM span task and the complex VSWM span task for younger children than for older children suggest at least an age-related differentiation into different components.

Notably, we did not test different WM models against each other (see, for example, Wilhelm, Hildebrandt, & Oberauer, 2013) and or data would not even allow for such an approach. We were interested in whether VWM and VSWM progressively differentiate with age, which we found to be consistent with prominent WM models (e.g., Baddeley, 2000; Barrouillet & Camos, 2012). Our main focus was the comparison of age-groups in childhood. Age-related differentiation and age-related performance differences in childhood are important characteristics of WM processing in specific and cognitive functioning in general, because WM capacity places limiting constraints for performing a wide range of other cognitive activities (e.g., Oberauer, 2009).

Older as well as younger children performed better in the VWM than in the VSWM span task. This is surprising given the hypothesis that especially younger children should be better in the VSWM span task, since the use of rehearsal strategies might not be fully developed. It was evident during testing that both age groups tended to name aloud the stimulus they saw on the screen in the VWM span task (cf. Karbach & Kray, 2007), indicating that younger children already used verbal rehearsal strategies. This is consistent with the fact that the younger sample in this study was already very verbally capable (see Table S1). Interestingly, children did not use overt verbal strategies in the VSWM span task. Moreover, the VWM and VSWM span task used in the present study were structurally and conceptually very similar, which may also have contributed to the fact that children did not perform worse on VWM span tasks.

The present study presents some limitations that need to be discussed. First, one might reason that the less pronounced differentiation of WM processes in younger children may be due to greater error measurement at that age. However, the correlations were tested in a multi-group model with strict measurement invariance. Thus, the intercepts and residuals for both groups were estimated on the same values, allowing for interpretation of the correlation between the factors. Second, we acknowledge that there is a slight difference between the two complex span tasks because of the processing task used during the encoding phase. Whereas the VWM task taps verbal memory and combines it with visual processing of the stimuli for the processing task (a cross-domain combination), the VSWM task taps spatial memory and visual processing only (a within-domain combination). Thus, interference between storage and processing may have been stronger in the VSWM task, which may have contributed to the finding that children performed better on the VWM than the VSWM task. However, this is just an assumption, since we did not include simple-span tasks in order to measure storage only.

In sum, we found age-related improvements in both VWM and VSWM from 4-6 to 8-10 years of age. We demonstrated measurement invariance across the age-groups, which is encouraging for future longitudinal research because it shows that these age groups can be compared on the same WM measures. At the same time, a differentiation into more separable VWM and VSWM components occurred, pointing to significant changes in the structure of WM during middle childhood. Further research with more working memory tasks is needed to further differentiate between different theoretical models and their subcomponents, and to understand how the system matures during normal development.

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Footnotes

¹ Ethical approval was given by the research ethics committee of the University of Edinburgh, protocol number 111-1516/2.

² One might criticize the fact that the two indicators for each WM domain were derived from the same task. Of course, two values derived from the same tasks should naturally correlate more strongly than different tasks. However, this has no influence on whether or not the four indicators (two VWM and two VSWM indicators) load on one common factor. Reliable tasks measuring the same underlying psychological construct have common variance over and above the task specific variance some indicators might share. Further, the question of one vs. two factors is not central here. The distinction between two strongly correlating factors and one common factor is a cut off on a continuous scale. Rather the notable different correlations between VWM and VSWM for the younger and older children are theoretically important.

Supplementary Material

This section includes:

Description of the control measures

Table S1. Descriptive statistics and reliabilities for the working-memory tasks, separated by age group

 Table S2. Correlation between performance on the working-memory tasks and control measures,

 separated by age group

 Table S3. Latent factor means for verbal and visuo-spatial working-memory performance

 Additional references

Description of the control measures

The *Peabody Picture Vocabulary Test* (PPVT-4, Dunn & Dunn, 2007) is a test for receptive language ability. Children were to point to one of four pictures that showed the referent of a word read out loud by the experimenter. We applied a short version of the test including a subset of 22 items, yielding a good reliability (split-half reliability: r = .85, odd-even split, Spearman-Brown corrected). The test score was the number of correct responses. Analysis of variance with the between-subjects factors Age group (younger, older) and Country (Germany, Scotland) showed a main effect of Age group, F(1, 222) = 278.78, $p \le .001$, $\eta p^2 = .56$, indicating that older children (M = 20.6, SE = 0.1) outperformed younger children (M = 15.0, SE = 0.3). Importantly, none of the remaining main effects and interactions reached significance (all *p*-values $\ge .334$).

The *Color-Naming Test* (Karbach, Kray, & Hommel, 2011) is a test for perceptual speed of processing. In this test, four different shapes are assigned to four specific colors (yellow circle, blue cross, red triangle, and green square). Forty-eight uncolored shapes were shown below the template (five practice items and 43 test items). After practice, children are asked to name the color for each shape as quickly and accurately as possible. The test showed excellent reliability (split-half reliability¹: r = .98, odd-even split, Spearman-Brown corrected). Children's score of processing speed was the sum of correctly named items within 45 seconds. Analysis of variance with the between-subjects factors Age group (younger, older) and Country (Germany, Scotland) showed a main effect of Age group, F(1, 222) = 262.05, $p \le .001$, $\eta p^2 = .54$), indicating that older children (M = 33.2, SE = 0.7) outperformed younger children (M = 17.4, SE = 0.7). Importantly, none of the remaining main effects and interactions reached significance (all *p*-values $\ge .366$).

Table S1

Descriptive statistics and reliabilities for the working-memory tasks as a function of age group

	Younger (4-6 year-olds, <i>n</i> = 125)			Older (8-10 year-olds, <i>n</i> = 101)			
	М	Maximum	Reliability	М	Maximum	Reliability	
	(SE)	(SE)		(SE)	(SE)		
Verbal WM	2.1 (0.05)	2.4 (0.09)	0.87	3.5 (0.06)	4.3 (0.08)	0.87	
Visuo-Spatial	1.8 (0.04)	2.1 (0.07)	0.82	3.4 (0.07)	4.3 (0.10)	0.90	
WM							

Note: Children's scores in the verbal and visuo-spatial WM span tasks. Children's receptive language ability was assessed with a subset (22 items) of the Peabody Picture Vocabulary test (PPVT-4, Dunn & Dunn, 2007). Processing speed was tested with the Color-Naming Test (Karbach, Kray, & Hommel, 2011). Reliabilities for the working-memory tasks were computed by relating the proportion of latent variation to total variation, which is equivalent to Cronbach's alpha (cf. Wilhelm & Schoebi, 2007). Numbers in brackets represent the standard error (SE).

¹ Due to technical problems, item-based information was available for the German sample. Thus, reliability analysis for the color-naming test was only based on N = 113 children.

Table S2

Correlation between performance on the working memory tasks and control measures as a

function of age group.

	Younger $(n = 125)$				Older ($n = 101$)			
Ability	1	2	3	4	1	2	3	4
1 Verbal WM	1	.57***	.47***	.40***	1	.44***	.38***	.27**
2 Visuo-Spatial WM	-	1	.28***	.37***	-	1	.31**	.29**
3 Receptive Language	-	-	1	.37***	-	-	1	.27**
4 Processing Speed	-	-	-	1	-	-	-	1

Note: p < .05, ** p < .01, *** $p \le .001$.

Table S3

Latent factor means for verbal and visuo-spatial working-memory performance in children

	Younger children	Older Children	
Verbal WM	$M_I = 2.07 (SD_I = 0.58)$	$M_2 = 3.50 \ (SD_2 = \ 0.57)$	$\Delta M_{2-1} = 1.43$, Cohen's $d = 13.96$
			$[\Delta \chi^2 (df = 1) = 176.23, p < .01]$
Visuo-Spatial	$M_3 = 1.80 \ (SD_3 = 0.42)$	$M_4 = 3.37 \ (SD_4 = 0.68)$	$\Delta M_{4-3} = 1.57$, Cohen's $d = 4.17$
			$[\Delta \chi^2 (df = 1) = 112.88, p < .01]$
	$\Delta M_{1-3}=0.27,$	$\Delta M_{2-4} = 0.13,$	
	Cohen's $d = 0.96$	Cohen's $d = 0.48$	
	$[\Delta \chi^2(df=1)=43.32,$	$[\Delta \chi^2(df=1)=4.21,$	
	<i>p</i> < .01]	<i>p</i> = .04]	

Notes. Mean group differences (ΔM) were estimated in a two-group CFA with strict group measurement invariance across age-groups (see Figure 1). The model-fit is displayed in Table 1 (line 4). Please note that the estimate of a latent mean is largely determined by the first indicator of a factor (here: the average span). It is closely correlated but not identical to the observed value of the first indicator. Cohen's *d* was calculated as the difference between two latent group means divided by their pooled standard deviation. Significance of ΔM was evaluated by SB-scaled χ^2 -Difference testing (i.e. comparing models with freely estimated means to models with means that were constrained to be equal across groups).

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