TITLE PAGE

Pre-pregnancy obesity and risk of congenital abnormalities of the kidney and urinary tract (CAKUT) – systematic review, meta-analysis and ecological study

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ABSTRACT:

INTRODUCTION: There is increasing evidence that maternal obesity is associated with several structural birth defects. Congenital abnormalities of the kidney and urinary tract (CAKUT) account for 30 to 50% of children starting kidney replacement therapy (KRT). We conducted a systematic review, meta-analysis and ecological study to explore the relationship between maternal obesity and CAKUT.

METHODS: A systematic literature search was conducted in EMBASE, MEDLINE, Global Health, The Cochrane Library, Scopus and Web of Science. Study quality was assessed for bias and confounding. A meta-analysis using a random effect model was carried out to obtain a summary odds ratio (OR) and 95% confidence interval (CI). In the ecological study, country level data were used to examine the correlation of secular trends in female obesity, CAKUT incidence and incidence of KRT.

RESULTS: Eight epidemiological studies were included in the review, 4 cohort studies and 4 case-control studies, 7 of which were included in the meta-analysis. There was evidence of a positive association between obesity during pregnancy and the risk of CAKUT, with a summary OR = 1.14 (1.02 - 1.27). No association was seen with overweight, nor a dose response with increasing obesity. There was an increasing trend in countries' proportion of female obesity and an increasing trend in reported CAKUT incidence with specific rises seen in congenital hydronephrosis (CH) and multi-cystic kidney dysplasia (MCKD).

CONCLUSION: Our findings suggest that pre-pregnancy obesity may be associated with increased risk of CAKUT at population level.

Key words: pre-pregnancy obesity; CAKUT (congenital anomalies of the kidney and urinary tract); meta-analysis; systematic review; ecological study

INTRODUCTION:

Obesity is a major public health issue and its prevalence continues to rise worldwide[1]. There is increasing evidence that maternal obesity is associated with several structural birth defects[2-4]. Congenital abnormalities of the kidney and urinary tract (CAKUT) account for 30 to 50% of children starting KRT[5] and occur in around 0.4 to 4 per 1000 pregnancies [6,7]. Although milder forms of CAKUT can remain asymptomatic, others will give rise either in infancy or later in life to recurrent urinary tract infections, hypertension or kidney insufficiency[6,8]. CAKUT is often an isolated condition with gene abnormalities accounting for approximately 15% of cases[9], can occur in families and in hundreds of syndromes[10]. Variations of the CAKUT phenotype within families and the variable expressivity of CAKUTcausing genes point to environmental factors, maternal obesity being one of them [9]. Animal studies have shown that maternal hyperglycaemia and maternal obesity can affect kidney development in the offspring[11,12]. Maternal type 1 diabetes mellitus (DM) and gestational DM (GDM) are known to increase the risk of congenital malformations[13,14] including CAKUT[15]. The effect of maternal obesity on CAKUT is less well known with studies showing different results[16,17].

The aim of this study was to systematically review the evidence and pool data that examined the association of obesity during pregnancy and CAKUT. As a secondary aim we conducted an ecological study to examine population trends in maternal obesity and incidence of CAKUT.

MATERIALS and METHODS

Systematic Review and Meta-Analysis:

These were conducted following PRISMA guidelines[18]. The protocol is registered on PROSPERO (PROSPERO ID 121709).

Criteria for inclusion were that studies were conducted in humans, had obesity as the defined exposure, indicated congenital abnormalities as an outcome, had a direct control group, were published in English and were peer-reviewed.

The literature search was conducted up to November 2018 using EMBASE (from 1947),
MEDLINE (from 1879), Global Health (from 1973), The Cochrane Library (from 1996), Scopus
(from 2004) and Web of Science (from 1900).

Searches using free-text and MeSH terms were done and the domains were obesity, pregnancy, congenital anomalies/abnormalities and birth registry or medical records, in order to identify studies on large populations with sufficient number of CAKUT outcomes (Supplementary File (SF) A). Under the domain "congenital abnormalities, we included: CAKUT, Vesico-ureteric/vesico-ureteral reflux, Renal/kidney dysplasia, Renal/kidney dysgenesis, Renal/kidney agenesis, Renal/kidney hypoplasia, Multicystic kidney dysplasia (MCD), Obstructive urinary defect, Obstructive genito-urinary defect, Megaureter, Pelvic ureteric, Hydronephrosis, Posterior urethral valves, Renal abnormality/anomaly, Urinary abnormality/anomaly, Kidney abnormality/anomaly, Renal/kidney cysts/cystic.

Study selection

Duplicates were removed. All titles and abstracts were examined by one reviewer (HA). Full text searches were made looking for urinary/kidney abnormalities if these were not mentioned in the abstract. A sample of 100 references were selected randomly and evaluated by a second reviewer (DN). There was full agreement.

Data collection process

A predesigned form was used to record extracted data that included study design, author name, year, setting, study population, CAKUT abnormality, definition of obesity and obesity subgroups, ascertainment, selection of control group, number of CAKUT cases in non-obese and in obese groups, crude and adjusted odds ratio (OR) and 95% confidence intervals (95%CI) and confounders.

Quality assessment

Risk of bias was assessed by HA, DN and LJ, using the following criteria: selection bias, information bias in the measurement of pre-pregnancy BMI and in the diagnosis and ascertainment of CAKUT cases. Study quality was further assessed by looking at adjustment for key confounders: maternal age, race/ethnicity, smoking, maternal diabetes status, folic acid use and socioeconomic status: maternal education, maternal parity (SF B for other desirable confounders). Studies were assigned high risk, low risk or uncertain according to the above criteria and were colour coded.

Meta-analysis

The main exposure was maternal obesity at time of conception or during the first trimester of pregnancy ("pre-pregnancy obesity"). The outcome was any form of CAKUT as defined in individual studies. Meta-analysis was performed in studies containing the OR or risk ratio (RR), comparing the incidence of CAKUT in maternal obese group and in non-obese group. If a study only reported a risk ratio for CAKUT for different levels of obesity, each measure of effect was included in the meta-analysis, otherwise we used the overall measure. The standard error for each study was calculated from log OR and log 95% CI. A pooled adjusted OR and 95% CI across studies was computed using a random effects model using Stata version 13 (www.stata.com). A subgroup analysis was carried based on study design. The

data were displayed in forest plots. We also obtained a pooled estimate for the association between CAKUT and obesity subclasses and overweight.

Heterogeneity was quantified by calculating I^2 , classified as 0-40% = might not be important; 30-60% = moderate; 50-90% = may represent substantial heterogeneity; 75-100% = considerable heterogeneity [19]. To exclude the possibility that a study exerted excessive influence on the heterogeneity, we conducted a sensitivity analysis by omitting 1 study at a time. We assessed the presence of publication bias using funnel plots and the Egger test where a p value of less than 0.10 was interpreted as supporting publication bias.

Ecological study

The main exposure for the ecological study was the countries' age standardised prevalence (ASP) of obesity per year in women aged 18 years or older. Obesity was defined as a BMI of 30 or above. We used the data published by the WHO (SFC). Other country level exposures explored were the ASP of raised fasting blood glucose and smoking in females, countries' gross domestic product (GDP), inequality using the GINI index. Outcome data consisted of a) the annual incidence per country of CAKUT (overall and specific diagnosis), data derived from EUROCAT[20], a European network of population-based registries for the surveillance of congenital abnormalities where cases are classified based on ICD10 with the British Paediatric Association extension and include livebirths, foetal deaths from 20 weeks gestational age and terminations of pregnancy for foetal anomaly at any gestation. The urinary abnormalities included in EUROCAT's surveillance are: bilateral kidneyagenesis, MCKD, congenital hydronephrosis (kidney pelvis diameter ≥ 10mm after birth), bladder exstrophy, posterior urethral valves and/or prune belly) b) the annual proportion per country of CAKUT as the primary kidney pathology in children under 15 years entering RRT in European countries derived from ERA-EDTA data and c) the incidence per country of young

people below 21, 18 or 16 years having RRT. Please seeSF C for RRT and other country level sources.

RESULTS

Systematic review

We initially identified 2107 articles. After removal of duplicates 1417 articles remained. Forty-five studies met our criteria and after assessment of their full text, 37 were excluded (Figure 1, SF D). The number of studies retrieved from each database is given in SF E. Eight studies were left after exclusions: 4 cohort studies (Block[17], Blomberg[21], Garcia-Patterson[22] and Persson[23]) and four case-control studies (Honein[24], Oddy[25], Slickers[16] and Tromp[26]), see Table 1.

Exposure Definition:

Obesity was defined using the WHO classification of BMI, as a BMI \geq 30 [27] in all studies except Garcia-Patterson which used BMI tertiles, upper tertile was a BMI of 24.78-47.07. This was similar to the "high BMI" group in the Honein study (BMI \geq 25). Three cohort studies (Block, Blomberg and Persson) further subdivided obesity into 3 WHO subclasses (Class I: BMI \geq 30 \leq 35, Class II: BMI \geq 35 \leq 40 and Class III: BMI \geq 40)[28]; Block and Persson provided effect measures for obesity and for the specific subclasses of obesity, Persson provided these for each subclass only. Except for the Honein study, the rest of the studies provided a separate effect measure for overweight; this corresponded to the 2nd tertile in the Garcia-Patterson study. All but one study used the pre-pregnancy weight as recollected by the mother for the pre-pregnancy BMI calculation. The Persson study used early pregnancy BMI using measured weight at the first antenatal visit in the first trimester.

Outcome Definition:

The diagnosis of CAKUT was based on the International Classification of Diseases, edition 9 or 10 in all studies except for the Garcia-Patterson which used a non-standardised definition. The outcome was reported with all CAKUT diagnoses grouped together under "Urinary Abnormalities" using ICD-9 or 10 codes in Persson and Oddy, without further details, excluding known chromosomal, genetic or syndrome diagnoses. Garcia-Patterson grouped them as Renal/Urinary without further details. The rest of the studies provided definition of outcomes giving details of the specific diagnoses included. Persson had a separate Genital abnormalities outcome. Garcia-Patterson and Oddy didn't and it is possible that minor anomalies, especially hypospadias, were included. Specific CAKUT diagnoses were used as outcome in Block (Obstructive Genito-Urinary Defect, Renal Agenesis (RA)/hypoplasia) and in Blomberg (Cystic kidney, RA/hypoplasia), in Honein (RA/Hypoplasia, Obstructive defects, Renal or Ureter duplication), in Slickers (RA/Hypoplasia), in Tromp (Bilateral RA, MCD, congenital hydronephrosis (kidney AP≥ 10 mm), bladder exstrophy or epispadias, posterior urethral valves or prune belly syndrome).

The age at diagnosis of CAKUT was at birth or in the neonatal period in Blomberg, Persson and Garcia-Patterson, by 1 year of age in Block, Honein and Slickers, by 6 years of age in Oddy and by 10 years of age in Tromp.

Quality Assessment:

Cohort studies: There were four cohort studies, two from Sweden (Blomberg and Persson) which were nationally based, one from Spain, based at a hospital in Barcelona (Garcia-Patterson) and a regional study in the USA (Block) based in Florida (Table 1). The total sample size consisted of 3419969 pregnancies and 9459 CAKUT cases. The populations in Block, Blomberg and Persson were over a million with hundreds of cases, in contrast to the

small cohort in Garcia-Patterson with just over two thousand participants. The study periods span the years from 1986 to 2014. Blomberg, Block and Persson excluded known chromosomal syndromes; in addition, Persson excluded infants with genetic or malformation syndromes as well as congenital viral infections linked to malformations. This is unknown for the Garcia-Patterson study.

Only Block excluded mothers with pre-existing diabetes but included GDM. Blomberg and Persson included both types. GDM was included in all studies and was present in 7,5% of obese mothers in Block, not stated in the others. Only Persson carried out a sensitivity analysis excluding all diabetes which showed no difference in the estimate. Garcia-Paterson studied high BMI in mothers with GDM.

Case-control studies: The four case-controlled studies originated from Australia (Oddy), the USA (Honein and Slickers) and from the Netherlands (Tromp) (Table 1). There was a total of 765 cases and 5616 controls. The oldest study is Honein with data from 1968 to 1980. The three other studies have data from 1997 to 2010. Cases were identified from regional or state level registries and included CAKUT occurring in live births, stillbirths and foetal deaths. Cases with known chromosomal abnormalities were excluded. Controls were selected from healthy infants born in the same area and period of time as the cases except for Tromp where controls were live infants or foetuses with a known syndrome, chromosomal or monogenic anomaly. Slickers and Honein excluded mothers with pre-existing diabetes. Oddy performed a sensitivity analysis which showed no difference. GDM was excluded by Honein and adjusted for by Slickers but not by Oddy or Tromp and was present in 1.2% of obese women in the Tromp study, not stated in the others.

See Table 2 for summary of quality criteria and assessment (details in SF F). Overall, the quality of cohort studies was higher showing low selection bias through the use of population-based

birth registries with the exception of Garcia-Patterson where the selection process was uncertain. There was a lack of detail on the conditions of the exposure measurement in Block and Garcia-Patterson. Measurement of the outcome was mostly unbiased helped using birth registries which had defined diagnostic classification and the use of defined age of ascertainment of the diagnosis of CAKUT, except for the Blomberg study where this varied from birth to 13 years of age. In all but one study there was an acceptable level of adjustment for confounders.

The case-control studies had a high risk of information bias arising from the use of self-reported maternal pre-pregnant BMI obtained after the outcome and at various times after birth. Adjustment for key confounders was poor in two and adequate in two studies. The effect estimate was higher in case-control studies, which also had wider confidence limits.

Meta-analysis

García-Patterson study was excluded from the meta-analysis because it was conducted amongst women with GDM, which is on the causal pathway between BMI (the exposure of interest) and CAKUT. This left 4 case control studies and 3 cohorts.

The individual study ORs were of similar order of magnitude with considerable overlap of 95% CI (Figure 2). The results indicated a positive association between pre-pregnancy obesity and CAKUT with an overall pooled OR of 1.14 (1.02 – 1.27). There was a moderate degree of heterogeneity with an I^2 of 51.5%, p=0.024, with the Persson study explaining most of the heterogeneity. When this study was omitted the OR was 1.23 (1.09-1.39) I^2 =26.7%, p=0.216. Omitting the Honein paper (which grouped together overweight and obese) from the obese group resulted in a very minimal increase in the summary estimate (see SF G). The subgroup cohort summary estimate was smaller and had narrower 95% CI

than the that of case-control studies although there was low statistical evidence of a difference (p=0.06). There was no clear evidence of a "dose-response" effect with increasing Obesity (SF H). There was no association with overweight, pooled OR was 1.02 (0.97 - 1.07), $I^2 = 0.0\%$, p=0.958 (SF I). Neither the funnel plot nor the Egger test suggested publication bias (SF J).

Ecological study:

Data from 21 European countries from 1980 show an upward trend in female obesity (Figure 3) which coincides with an increase in the yearly incidence of CAKUT (Figure 4). When the relationship with female obesity is examined in more detail for specific CAKUT diagnoses, only the incidence of Congenital Hydronephrosis and of MCKD seem to increase whilst the other two remain stable[20] (SF K). Individual country plots of proportion of female obesity against each specific CAKUT diagnosis seemed to confirm this pattern (SF L). There was no increase in the proportion of children with a CAKUT diagnosis entering KRT when looking at European data from 1980 onwards (SF M-a). We then explored whether secular trends in children starting KRT mirrored the increase in female obesity. We assumed that the proportion entering KRT with a CAKUT diagnosis would remain the same in each country over the period studied. Using data from a wide range of countries worldwide on the yearly incidence of all children starting KRT from the late 1990s we found that there was no overall proportional increase (SF M-b). The only country where the RRT incidence (in < 25-year olds) mirrored the increase in female obesity was Malaysia. There were no associations between ASP of female smoking and glycaemia, countries' GDP or Gini index (SF N

DISCUSSION

In this study, we have We found that maternal obesity may be a risk factor for CAKUT. The meta-analysis showed a small but positive association between pre-pregnancy obesity and CAKUT. This was not seen with overweight in pregnancy nor did we find a dose response with increasing obesity. The size of the association is similar to the level of risk found in comparable meta-analysis of pre-pregnancy obesity and other congenital defects[4]. The results of the ecological study suggested that as female obesity increased so had the overall number of CAKUT diagnoses.

The mechanisms for this association are not clear. Maternal obesity is associated with a wide range of metabolic abnormalities, but little is known about their potential effect epigenetic effects.[29]

Our results were mainly driven by papers based on large-population birth registries with most studies using external, well standardised definitions of the exposure and outcome. Cohorts and case-control studies were characterised by low selection bias since data were mainly obtained from birth registries and by low bias in assessing the outcome. Most studies had adjusted for key confounders, and therefore this small association seen may be potentially real.

Our study suffers from some limitations. We conducted the meta-analysis on aggregated data rather than individual data. However, the absolute number of participants and events was very high and therefore the summary estimate is likely to be similar with either method[30,31]. In the studies included in the meta-analysis, there was clinical heterogeneity arising from differences in outcomes definitions, with CAKUT diagnoses either grouped together or analysed individually. Grouping may have increased the power of these studies but this may not be the right approach as the risk may vary depending on the specific

CAKUT diagnosis. There was some suggestion from the ecological study that only congenital hydronephrosis and MCKD showed increases as female obesity increased.. Confounding by pre-existing diabetes was adequately addressed except for one study (Bloomberg). All studies in the meta-analysis except Honein's included infants of mothers with GDM which may lie in the causal pathway between obesity and CAKUT. Only Persson performed a sensitivity analysis excluding GDM mothers. Slickers adjusted for GDM; however, if GDM lies in the causal pathway, this adjustment would probably result in a lower estimate. However, where available, the proportion of obese women with GDM was very small. The association of gestational DM and CAKUT was independent of pre-pregnancy BMI[15]. Although our results would suggest that pre-pregnancy BMI may have a threshold effect on kidney development, the absence of a dose related effect might be due to small numbers of CAKUT outcomes in the obesity sub-categories. The ecological study may be showing secular increases in the ascertainment of CAKUT related to increased reporting, for example of mild congenital hydronephrosis, or to the greater availability of prenatal ultrasound, occurring in parallel to increases in female obesity. This explanation, however, is less likely to apply to conditions such as MCKD, which shows definitive and more obvious features on prenatal ultrasound. . The findings from the ecological study showing secular, parallel increases in female obesity and in the diagnoses of congenital hydronephrosis and MCKD should not be used to interpret individual risk and should instead be regarded as sources for hypotheses to be tested in further studies. Diagnosis evident at birth rely on prenatal ultrasound findings, usually performed by 20 weeks gestation. CAKUT may become apparent later in gestation and ultrasonography is technically more difficult in obesity, conceivably missing a CAKUT diagnosis altogether. CAKUT may become apparent later in life and may be underdiagnosed in young adults presenting in CKD stage 5[32]. This differential

bias would underestimate the effect of maternal obesity. Case-control studies had wider 95% CI and their use of self-reported pre-pregnancy weight increased information bias. The cohort sub-estimate, although smaller, may be more accurate. There is a lack of studies from middle- and low-income countries.

Implications:

This increase in risk, even though small, needs to be seen as potentially adding to the burden of disease linked to pre-pregnancy obesity such as increased risk of other congenital abnormalities as already mentioned and increased risk in long term obesity and increased cardiometabolic risk in the offspring[33,34]. There is a public health dimension to the findings of our study because obesity continues to rise worldwide, especially in females[1]. Management of obesity at the start of antenatal care may be too late because the first glomeruli form between 9 to 10 weeks gestation and the kidney development will have completed its morphological changes by 20 weeks. Public health measures that address the obesogenic nature of living environments are going to be more effective.

The association found is small and it is not possible at this stage to be certain about the clinical implications of our findings on an individual context beyond continuing being proactive on encouraging a healthy lifestyle in our paediatric patients.

This association should be the subject of further studies in different populations, focussing on specific CAKUT diagnoses.. The importance of birth registries in this context is paramount. The impact of the associations seen may be greater in settings where there is poor or no access to paediatric urological care. Further research on potential mechanisms to explain this association is needed.

Conclusions

Our analysis suggests that pre-pregnancy obesity may be a risk factor for CAKUT, but this

finding needs to be explored further within specific CAKUT diagnoses.

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REFERENCES

- 1. WHO Obesity and Overweight. https://www.who.int/news-room/fact-sheets/detail/obesity-and-overweight. Accessed 26/02/2020 2020
- 2. Stothard KJ, Tennant PW, Bell R, Rankin J (2009) Maternal overweight and obesity and the risk of congenital anomalies: a systematic review and meta-analysis. JAMA 301 (6):636-650. doi:10.1001/jama.2009.113
- 3. Huang HY, Chen HL, Feng LP (2017) Maternal obesity and the risk of neural tube defects in offspring: A meta-analysis. Obes Res Clin Pract 11 (2):188-197. doi:10.1016/j.orcp.2016.04.005
- 4. Zhu Y, Chen Y, Feng Y, Yu D, Mo X (2018) Association between maternal body mass index and congenital heart defects in infants: A meta-analysis. Congenit Heart Dis 13 (2):271-281. doi:10.1111/chd.12567
- 5. Harambat J, van Stralen KJ, Kim JJ, Tizard EJ (2012) Epidemiology of chronic kidney disease in children. Pediatr Nephrol 27 (3):363-373. doi:10.1007/s00467-011-1939-1
- 6. Andres-Jensen L, Jorgensen FS, Thorup J, Flachs J, Madsen JL, Maroun LL et al. (2016) The outcome of antenatal ultrasound diagnosed anomalies of the kidney and urinary tract in a large Danish birth cohort. Arch Dis Child 101 (9):819-824. doi:10.1136/archdischild-2015-309784
- 7. Tain Y-L, Luh H, Lin C-Y, Hsu C-N (2016) Incidence and Risks of Congenital Anomalies of Kidney and Urinary Tract in Newborns: A Population-Based Case-Control Study in Taiwan. Medicine 95 (5):e2659-e2659. doi:10.1097/MD.0000000000002659
- 8. Nef S, Neuhaus TJ, Sparta G, Weitz M, Buder K, Wisser J et al. (2016) Outcome after prenatal diagnosis of congenital anomalies of the kidney and urinary tract. Eur J Pediatr 175 (5):667-676. doi:10.1007/s00431-015-2687-1
- 9. Nicolaou N, Renkema KY, Bongers EM, Giles RH, Knoers NV (2015) Genetic, environmental, and epigenetic factors involved in CAKUT. Nat Rev Nephrol 11 (12):720-731. doi:10.1038/nrneph.2015.140
- 10. Limwongse C (2009) Syndromes and Malformations of the Urinary Tract. In: Editors: Avner EDHW HW, Niaudet P, Yoshikawa N (ed) Pediatric Nephrology, 6th Edition, vol 1. Springer, Berlin, pp 122-138
- 11. Kanwar YS, Nayak B, Lin S, Akagi S, Xie P, Wada J et al. (2005) Hyperglycemia: its imminent effects on mammalian nephrogenesis. Pediatr Nephrol 20 (7):858-866. doi:10.1007/s00467-005-1888-7 12. Glastras SJ, Chen H, McGrath RT, Zaky AA, Gill AJ, Pollock CA et al. (2016) Effect of GLP-1 Receptor Activation on Offspring Kidney Health in a Rat Model of Maternal Obesity. Sci Rep 6:23525-23525. doi:10.1038/srep23525

- 13. Zhao E, Zhang Y, Zeng X, Liu B (2015) Association between maternal diabetes mellitus and the risk of congenital malformations: A meta-analysis of cohort studies. Drug Discov Ther 9 (4):274-281. doi:10.5582/ddt.2015.01044
- 14. Balsells M, Garcia-Patterson A, Gich I, Corcoy R (2009) Maternal and fetal outcome in women with type 2 versus type 1 diabetes mellitus: a systematic review and metaanalysis. J Clin Endocrinol Metab 94 (11):4284-4291. doi:10.1210/jc.2009-1231
- 15. Parimi M, Nitsch D (2020) A Systematic Review and Meta-Analysis of Diabetes During Pregnancy and Congenital Genitourinary Abnormalities. Kidney International Reports. doi:https://doi.org/10.1016/j.ekir.2020.02.1027
- 16. Slickers JE, Olshan AF, Siega-Riz AM, Honein MA, Aylsworth AS (2008) Maternal body mass index and lifestyle exposures and the risk of bilateral renal agenesis or hypoplasia: the National Birth Defects Prevention Study. Am J Epidemiol 168 (11):1259-1267. doi:10.1093/aje/kwn248
- 17. Block SR, Watkins SM, Salemi JL, Rutkowski R, Tanner JP, Correia JA et al. (2013) Maternal prepregnancy body mass index and risk of selected birth defects: evidence of a dose-response relationship. Paediatr Perinat Epidemiol 27 (6):521-531. doi:10.1111/ppe.12084
- 18. Moher D, Liberati A, Tetzlaff J, Altman DG (2009) Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. BMJ 339. doi:10.1136/bmj.b2535
- 19. Higgins JPT G, S. Cochrane Handbook for Systematic Reviews of Interventions Version 5.1.0, Chapter 10. The Cochrane Collaboration; 2011, .

https://training.cochrane.org/handbook/current/chapter-10.

- 20. EUROCAT European network of population-based registries for the epidemiological surveillance of congenital anomalies. European Commission. https://eu-rd-platform.jrc.ec.europa.eu/eurocat. 2019
- 21. Blomberg MI, Kallen B (2010) Maternal obesity and morbid obesity: the risk for birth defects in the offspring. Birth Defects Res A Clin Mol Teratol 88 (1):35-40. doi:10.1002/bdra.20620
- 22. Garcia-Patterson A, Erdozain L, Ginovart G, Adelantado JM, Cubero JM, Gallo G et al. (2004) In human gestational diabetes mellitus congenital malformations are related to pre-pregnancy body mass index and to severity of diabetes. Diabetologia 47 (3):509-514. doi:10.1007/s00125-004-1337-3
- 23. Persson M, Cnattingius S, Villamor E, Soderling J, Pasternak B, Stephansson O et al. (2017) Risk of major congenital malformations in relation to maternal overweight and obesity severity: cohort study of 1.2 million singletons. BMJ 357:j2563. doi:10.1136/bmj.j2563
- 24. Honein MA, Moore CA, Watkins ML (2003) Subfertility and prepregnancy overweight/obesity: possible interaction between these risk factors in the etiology of congenital renal anomalies. Birth Defects Res A Clin Mol Teratol 67 (8):572-577. doi:10.1002/bdra.10077
- 25. Oddy WH, De Klerk NH, Miller M, Payne J, Bower C (2009) Association of maternal pre-pregnancy weight with birth defects: evidence from a case-control study in Western Australia. Aust N Z J Obstet Gynaecol 49 (1):11-15. doi:10.1111/j.1479-828X.2008.00934.x
- 26. Tromp L, de Walle HEK (2012) Pre-pregnant maternal obesity increases risk of several congenital anomalies, in particular neural tune defects in the offspring. UMCG EUROCAT Northern Netherlands

http://scripties.umcg.eldoc.ub.rug.nl/FILES/root/geneeskunde/2012/TrompL/TrompL.pdf

- 27. WHO: World health statistics 2009. Geneva: World Health Organization 2009.
- 28. WHO (2000) Obesity: preventing and managing the global epidemic. Report of a WHO consultation. World Health Organ Tech Rep Ser 894:i-xii, 1-253
- 29. Catalano PM, Shankar K (2017) Obesity and pregnancy: mechanisms of short term and long term adverse consequences for mother and child. BMJ 356:j1. doi:10.1136/bmj.j1
- 30. Tudur Smith C, Marcucci M, Nolan SJ, Iorio A, Sudell M, Riley R et al. (2016) Individual participant data meta-analyses compared with meta-analyses based on aggregate data. Cochrane Database Syst Rev 9 (9):Mr000007. doi:10.1002/14651858.MR000007.pub3
- 31. Tierney JF, Fisher DJ, Burdett S, Stewart LA, Parmar MKB (2020) Comparison of aggregate and individual participant data approaches to meta-analysis of randomised trials: An observational study. PLoS Med 17 (1):e1003019. doi:10.1371/journal.pmed.1003019

- 32. Neild GH (2009) What do we know about chronic renal failure in young adults? I. Primary renal disease. Pediatr Nephrol 24 (10):1913-1919. doi:10.1007/s00467-008-1108-3
- 33. Gaillard R, Steegers EA, Duijts L, Felix JF, Hofman A, Franco OH et al. (2014) Childhood cardiometabolic outcomes of maternal obesity during pregnancy: the Generation R Study. Hypertension 63 (4):683-691. doi:10.1161/hypertensionaha.113.02671
- 34. Voerman E, Santos S, Patro Golab B, Amiano P, Ballester F, Barros H et al. (2019) Maternal body mass index, gestational weight gain, and the risk of overweight and obesity across childhood: An individual participant data meta-analysis. PLoS Med 16 (2):e1002744. doi:10.1371/journal.pmed.1002744
- 35. Organization WH Body Mass Index. http://www.euro.who.int/en/health-topics/disease-prevention/nutrition/a-healthy-lifestyle/body-mass-index-bmi.
- 36. Clinical Guidelines on the Identification, Evaluation, and Treatment of Overweight and Obesity in Adults--The Evidence Report. National Institutes of Health. Obesity research 1998, 6 Suppl 2:51S—209S.

Table 1. Characteristics of eligible studies

| | | | | | _ | | | | |
|----------|--------------|----------------------|---------------------|-----------------------|--------------------|----------------|--------------------------------------|-----------------|--|
| | | | Obesity | | CAKUT | | | Confounders | |
| | | | | | | | | | |
| Study | Setting | Population | Definition | Ascertained | Definition | Specification | Ascertained | | |
| Block | Florida, USA | Florida resident | WHO classification | FBDR (Florida Birth | Obstructive | As per | FBDR, diagnosed within 1st year of | Maternal age, | |
| 2013 | | mothers without | based on BMI[35]: | Defects Registry) – | genito-urinary | Definition | life, excluding trisomies 21, 18 and | race/ethnicity, | |
| 1013 | | pre-gestational | Underweight: <18.5 | modified to capture | defect and renal | column | 13 | education, | |
| | | diabetes who gave | onderweight. (10.5 | maternal pre- | agenesis/hypopla | | | smoking, marita | |
| | | birth to live | Normal: 18.5-24.9 | pregnancy weight | sia based on ICD-9 | | | status, and | |
| | | singleton infants | Overweight: 25.0- | and height in 2014 | codes | | | nativity. | |
| | | from March 2004 | 29.9 | | | | | | |
| | | through December | Obese I: 30.0-34.9 | | | | | | |
| | | 2009 (1,124,370 | Obese 1. 30.0-34.3 | | | | | | |
| | | infants, 3987 cases | Obese II: 35.0-39.9 | | | | | | |
| | | of Obstructive | Obese III: ≥ 40.0 | | | | | | |
| | | Genito-urinary | | | | | | | |
| | | Defect, 435 cases of | | | | | | | |
| | | renal agenesis | | | | | | | |
| | | /hypoplasia) | | | | | | | |
| | | | | | | | | | |
| Blomberg | Sweden | Mothers resident in | WHO classification | Swedish medical | ICD-9 | Kidney | Swedish Medical Birth Registry and | Maternal age, | |
| 2010 | | Sweden giving birth | based on BMI | birth register (based | codes 740.0- | agenesis/dysge | the Swedish Register of Birth | parity, smoking | |
| | | to live infants or | | on standardised | 759.9 or ICD-10 | | Defects both based on data from | | |

| | | stillbirth (after 28 | including Obesity | medical record | codes beginning | nesis, cystic | the neonatal period including | early pregnancy, |
|-----------|------------|--------------------------|-------------------|---------------------|--------------------|---------------|--------------------------------------|---------------------|
| | | weeks gestation) | classes I to III | forms completed at | with Q | kidney | stillbirths from 28 weeks gestation | and year of birth |
| | | between January 1, | | the maternity | | | and from the National Patient | |
| | | 1995, through | | health care centres | | • | Register which contains | |
| | | December 31, 2007 | | at the start of | | | information on discharge diagnoses | |
| | | with Diabetes | | prenatal care, | | | of all patients admitted to Swedish | |
| | | included (1,049,582 | | usually in | | | hospitals and was used up to and | |
| | | infants, 291 cases | | gestational week | | | including 2007. Age at diagnosis | |
| | | of kidney | | 10–12) | | | from birth to up to 13 years of age. | |
| | | agenesis/dysgenesis | | | | | It excluded infants with known | |
| | | 391 cases of cystic | | | | | chromosomal abnormalities | |
| | | kidney) | | | | | | |
| García- | Barcelona, | mothers with a | BMI grouped into | Not mentioned | Neonatologist | Renal/Urinary | Not mentioned, possibly from | Previous |
| Patterson | Spain | documented | tertiles: | | assessment | | Hospital records before discharge | hyperglycaemia, |
| 2004 | Spani | diagnosis of | 15.43-21.91 | | before | • | after birth | parameters |
| | | Gestational | | | hospital | | | related to severity |
| | | Diabetes Mellitus | 21.92-24.77 | | discharge, | | | of GDM, pre- |
| | | who gave birth <u>to</u> | 24.78-47.07 | | classified into | | | pregnancy BMI, |
| | | <u>liveborn</u> between | | | Major Congenital | | | maternal age, |
| | | January 1986 and | | | Malformation if | | | maternal smoking |
| | | July 2002 at ≥22 | | | caused significant | | | |
| | | complete weeks of | | | functional or | | | |
| | | gestation at a single | | | cosmetic | | | |
| | | institution (Hospital | | | impairment or | | | |
| | | of the Holy Cross | | | required surgery | | | |
| | | and St Paul, | | | or Minor if not | | | |

| | 1 | 1 | I | 1 | I | T | | T |
|---------|--------|---------------------|--------------------|-----------------------|--------------------|---------------|--------------------------------------|---------------------|
| | | Barcelona, Spain), | | | serious | | | |
| | | 2060 infants, | | | medical/cosmetic | | | |
| | | number of cases not | | | significance and | | | |
| | | provided | | | occurring in <4% | | | |
| | | | | | of the background | | | |
| | | | | | population. | | | |
| | | | | | | | | |
| Persson | Sweden | Mothers resident in | WHO classification | Swedish medical | ICD-10 | Urinary | Swedish medical birth register | Maternal age (13- |
| 2017 | | Sweden giving birth | based on BMI | birth register, based | classification and | malformations | using standardised prenatal, | 24, 25-29, 30-34, |
| | | to liveborn | including Obesity | on measured | were defined | | obstetric, and neonatal records, the | ≥35 years), height |
| | | singleton infants | classes I to III | weight and self- | according to the | | national patient register, cause of | (130-154, 155- |
| | | from 2001 to 2014 | | reported height in | European | | death register and total population | 159, 160-164, |
| | | with diabetes | | first trimester at | Surveillance of | | register which includes diagnoses | 165-169, 170-174, |
| | | included (1,243,957 | | health care centres | Congenital | | and dates on visits for hospital | 175-200 cm), |
| | | infants, 4211 cases | | | Anomalies | | based inpatient and outpatient | parity |
| | | of Urinary | | | classification | | care, diagnosed by 1 year of age. | (primiparous, |
| | | malformations) | | | Classification | | Exclusions: infants with | multiparous), |
| | | | | | | | chromosomal aberrations, genetic | early pregnancy |
| | | | | | | | syndromes, malformation | smoking status |
| | | | | | | | syndromes with known causes, and | (non-smoker, 1-9, |
| | | | | | | | viral infections having a possible | ≥10 cigarettes |
| | | | | | | | association with malformations. | daily), educational |
| | | | | | | | | level (<10, 10-12, |
| | | | | | | | | >12 years), |
| | | | | | | | | maternal country |
| | | | | | | | | of birth (Nordic |
| | | | | | | | | (Sweden, |
| | | | | 1 | 1 | | | |

| | | | | | | | | | Denmark, Finland, Iceland, and Norway), non- Nordic), family situation (living with partner, not living with |
|----------|--------------|---------------------|------------------------|-------------------|-------------------|-----------------|---------------|--------------------|--|
| | | | | | | | | | partner), and sex |
| | | | | | | | | | of offspring |
| Casecont | rol studies: | 1 | 1 | 1 | 1 | 1 | 1 | | 1 |
| | | | Obesity | | CAKUT | | | Control | Confounders |
| | | | | | | | | | adjusted |
| Study | Setting | Population | Definition | Ascertained | Definition | Specification | Ascertained | Selection | |
| Honein | Atlanta, | Infants born in the | National Institutes of | Interviewed by | Hospital | Congenital | Identified by | Control-infants | gravidity, sex, |
| 2003 | USA | five-county | Health[36] BMI | telephone in 1982 | Adaptation of the | kidney | the Centres | were randomly | birth weight, |
| | | metropolitan | classification then | and 1983 (pre- | International | anomalies: | for Disease | selected from live | gestational age, |
| | | Atlanta area during | grouped as: | pregnancy weight | Classification of | agenesis/hypopl | Control and | births | maternal |
| | | 1968 –1980, | high BMI = ≥ 25 | and height), BMI | Disease (H-ICDA) | asia, | Prevention's | to residents of | education, |
| | | excluded if mother | myn bivn – 2 23 | calculated by | or ICD-9 | Obstructive | population- | metropolitan | maternal age, and |
| | | had diabetes | low BMI = < 25 | clinician | classification | defects, kidney | based | Atlanta, and were | maternal smoking |
| | | Cases = 169 | | | | or ureter | surveillance | frequency- | |
| | | Cuses - 109 | | | | duplication. | system and | matched to case- | |
| | | Controls = 2763 | | | | | the | infants according | |
| | | | | | | | Metropolitan | to birth year, | |
| | | | | | | | Atlanta | | |

| California, lowa, Massachusetts, New Jersey, New York, Texas and Atlanta (liveborn or | lickers USA 008 | race, and birth hospital Randomly Study centre, selected at each maternal study education (<. | |
|---|--------------------|--|--------------------------|
| liveborn and foetal deaths depending on state), excluded if mother had pre- pregnancy Diabetes Cases = 75 Controls = 868 | | white, non- Hispanic blace Hispanic, oth maternal age (years), the presence or | nic ack, ther), ge dism, |

| | | | | | | | | | sought assistance |
|------|--------------|-----------------------|--------------------|----------------|---------------------|--------------|---------------|--------------------|-------------------|
| | | | | | | | | | for pregnancy |
| | | | | | | | | | through |
| | | | | | | | | | medications or |
| | | | | | | | | | procedures, late |
| | | | | | | | | | pregnancy |
| | | | | | | | | | identification |
| | | | | | | | | | (after 12 weeks), |
| | | | | | | | | | vasoactive |
| | | | | | | | | | substance use |
| | | | | | | | | | during the first |
| | | | | | | | | | trimester, and a |
| | | | | | | | | | continuous |
| | | | | | | | | | representation of |
| | | | | | | | | | the number of |
| | | | | | | | | | months that folic |
| | | | | | | | | | acid was used |
| | | | | | | | | | during this 4- |
| | | | | | | | | | month |
| | | | | | | | | | periconceptional |
| | | | | | | | | | period |
| Oddy | Australia | Infants (live births, | WHO classification | Self-reporting | Diagnoses based | Urinary | Western | A random sample | Marital status, |
| 2009 | , tasti alla | stillbirths and | based on BMI | questionnaires | on ICD-9 with | malformation | Australian | of all live born | maternal age, |
| 2003 | | terminations of | Dasca on bivii | questionnunes | British Association | manormation | Birth Defects | infants in Western | maternal |
| | | pregnancy) | | | extension | | Registry | Australia using | education and |
| | | between September | | | Malformation | | negistry | the | periconceptional |
| | | between September | | | wajomation | | | uit | periconceptional |

| | | 1997 and March | | | ascertained by | | | statutory | folic acid |
|-------|-------------|---------------------|--------------------|------------------|--------------------|------------------|-------------|-------------------|--------------------|
| | | 2000. Cases = 87 | | | four months after | | | Midwives' | supplementation. |
| | | Controls = 418 | | | birth/termination, | | | Notification | |
| | | Controls 110 | | | excluding cases | | | System of all | |
| | | | | | with a known | | | births in | |
| | | | | | chromosomal, | | | Western Australia | |
| | | | | | genetic or | | | as a sampling | |
| | | | | | syndrome | | | frame | |
| | | | | | | | | | |
| | | | | | | | | | |
| Tromp | Northern | Children and | WHO classification | EUROCAT Northern | ICD10 with the | Urinary | EUROCAT | Children and | Maternal age at |
| 2012 | Netherlands | foetuses born | based on BMI | Netherlands | British Paediatric | abnormalities: | Northern | foetuses born | delivery, smoking, |
| | | between 1997 and | including Obesity | database | Association (BPA) | Bilateral kidney | Netherlands | between 1997 | education, folic |
| | | 2010 and included | classes I to III | | one-digit | agenesis, | database | and 2010 with a | acid use and |
| | | live births, foetal | | | extension | multicystic | | syndrome, | history of |
| | | death, stillbirths, | | | | kidney | | chromosomal or | pregnancy |
| | | miscarriages and | | | | dysplasia, | | monogenic | affected by |
| | | elective | | | | congenital | | anomaly | congenital |
| | | terminations, no | | | | hydronephrosis | | | anomalies |
| | | exclusions related | | | | (kidney AP≥ 10 | | | |
| | | to maternal pre- | | | | mm), bladder | | | |
| | | pregnancy or | | | | exstophy or | | | |
| | | gestational | | | | epispadias, | | | |
| | | Diabetes | | | | posterior | | | |
| | | Cases 434 | | | | urethral valves | | | |
| | | Controls 1567 | | | | or prune belly | | | |
| | | | | | | syndrome | | | |
| | | | | | | | | | |

Table 2: Summary of quality assessment: green = low risk, yellow = uncertain risk, red = high risk of bias or confounding

| | Selection | bias | Information Bias (Exposure | Information Bias (Out | Confounding | | | | | |
|----------------|-------------------------|----------------------|--|-----------------------|---------------|---|-------------|------------------|------------------------|-------------|
| Study | Loss to fo | llow up | Non-differential misclassification of exposure | Recall bias | Observer bias | Non-differential misclassification of outcome | Recall bias | Observer bias | Ascertainmen t bias | Confounding |
| Block | | | | | | | | | | |
| Blomberg | | | | | | | | | | |
| García- | | | | | | | | | | |
| Patterson | | | | | | | | | | |
| Persson | | | | | | | | | | |
| Case-control s | tudies | | | | | | | | | |
| Selection Bias | | | Information Bias (Exposure) | | | Information Bias (Out | Confounding | | | |
| Study | Loss to follow up | Selection of control | Non-differential misclassification of exposure | Recall bias | Observer bias | Non-differential misclassification of outcome | Recall bias | Observer bias | Ascertainmen t bias | Confounding |
| Honein | | | | | | | | | | |
| Oddy | | | | | | | | | | |
| Slickers | | | | | | | | | | |

Tromp

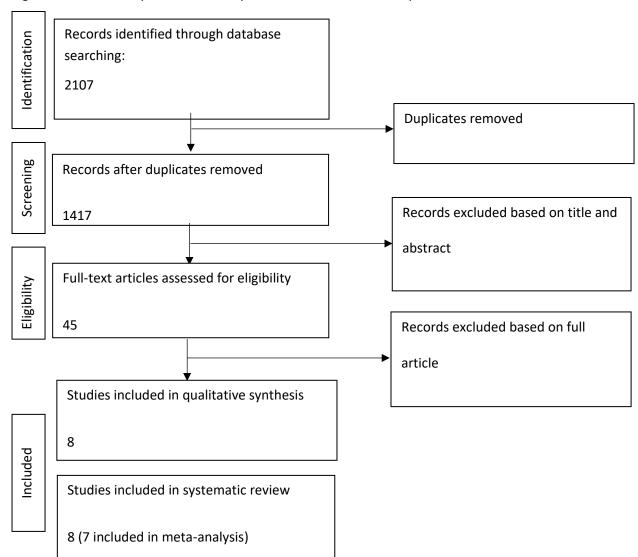
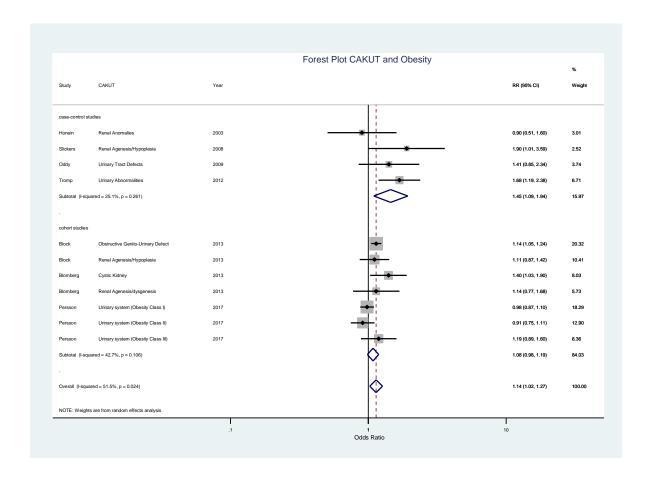
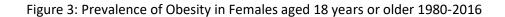


Figure 1. Flow of study selection into systematic review/meta-analysis

Figure 2: Forest Plot CAKUT and pre-pregnancy obesity





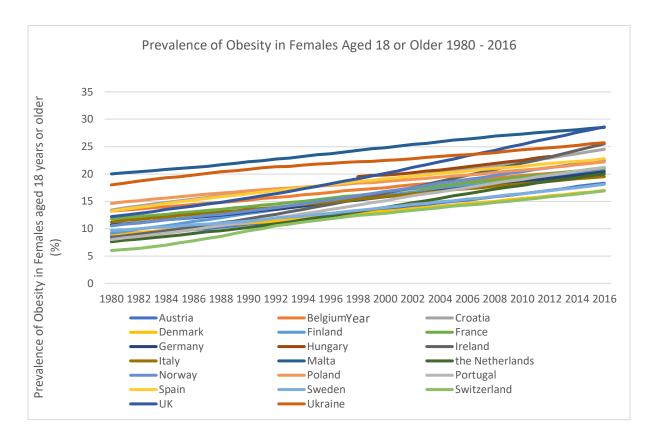


Figure 4: Incidence of CAKUT* in European countries 1980 to 2017 (anomalies included in CAKUT: congenital hydronephrosis, posterior urethral valves, prune-belly syndrome, multicystic kidney dysplasia and bilateral kidney agenesis)

