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1 **A theoretical framework for transitioning from patient-level to population-scale**
2 **epidemiological dynamics: influenza A as a case study**

3
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24 **ABSTRACT**

25 Multi-scale epidemic forecasting models have been used to inform population-scale
26 predictions with within-host models and/or infection data collected in longitudinal cohort
27 studies. However, most multi-scale models are complex and require significant modelling
28 expertise to run. We formulate an alternative multi-scale modelling framework using a
29 compartmental model with multiple infected stages. In the large-compartment limit, our
30 easy-to-use framework generates identical results compared to previous more
31 complicated approaches. We apply our framework to the case study of influenza A in
32 humans. By using a viral dynamics model to generate synthetic patient-level data, we
33 explore the effects of limited and inaccurate patient data on the accuracy of population-
34 scale forecasts. If infection data are collected daily, we find that a cohort of at least 40
35 patients is required for a mean population-scale forecasting error below 10%. Forecasting
36 errors may be reduced by including more patients in future cohort studies or by increasing
37 the frequency of observations for each patient. Our work therefore provides not only an
38 accessible epidemiological modelling framework, but also insight into the data required
39 for accurate forecasting using multi-scale models.

40 **KEYWORDS**

41 epidemiological model; infectious disease outbreak forecasting; multi-scale model;
42 nested model; longitudinal study; cohort study

43

44

1. INTRODUCTION

45 Infectious disease epidemics in humans, animals and plants have severe impacts [1–7].
46 Mathematical models are increasingly used to forecast the future dynamics of outbreaks

47 [7–9] and to plan interventions [10–13], while within-host models are used to understand
48 the spread of infection at the individual host-level [14–17]. Standard population-scale
49 epidemiological models assume that the infectiousness of each host is constant over
50 the course of the infectious period [4], but in reality infectiousness will vary as the
51 infection progresses through the host due to changing pathogen loads [18,19] and other
52 factors including behavioural responses to infection [18,20].

53

54 Multi-scale models have been used to connect epidemiological dynamics at the patient-
55 level (within-host; we refer to “patients” throughout but similar ideas can be applied to
56 pathogens of animals or plants) to those at the population-scale (between-host) [21–34].
57 These models (sometimes referred to as nested models [19,30]) tend to assume a
58 specified relationship between the level of infection within a patient and the rate at
59 which the patient transmits the pathogen to susceptible individuals [18,19,35]. A within-
60 host model, parameterised by fitting to patient data, is then used to determine the
61 parameters of a population-scale model incorporating time-dependent infectiousness
62 [19,35]. In addition to patient-level dynamics affecting population-scale transmission,
63 there may be reciprocal feedback from the population-scale to the patient-level [19] – for
64 example, if there are multiple co-circulating strains of the pathogen [24].

65

66 A recent review concluded that, while numerous multi-scale epidemiological modelling
67 studies exist, relatively few include substantial use of data [29]. While one reason for
68 this is the lack of widely available datasets [18,36], we contend that another contributing
69 factor is that previous multi-scale modelling frameworks have been complex, making

70 them challenging to implement other than by highly specialist mathematical modellers.
71 Such frameworks have often employed integro-differential equations (IDEs)
72 [19,24,27,30,31,33,35,37], although alternatives such as individual-based stochastic
73 models [12,23,25,26,38] have also been considered. IDEs are challenging to solve,
74 requiring bespoke numerical methods [28]. Some studies using IDEs have involved
75 explicit simulation of the full multi-scale model [24,31,37]. However, others have either
76 only used the multi-scale framework to derive quantities such as the basic reproduction
77 number of the pathogen rather than predicting temporal epidemic dynamics [30,33], or
78 have made simplifying assumptions such as taking a within-host model to be in
79 equilibrium [21,22,27,32]. Although an assumption that the pathogen load in each
80 infected host is not changing (or changes only a limited number of times) might be
81 appropriate for chronic infections, it leads to an approximate population-scale model
82 that does not explicitly account for time-dependent infectiousness or other potentially
83 complex patient-level dynamics.

84

85 In most previous studies that have used IDEs to transition from within- to between-host,
86 the progression of infection through all patients has been assumed to be identical [35].
87 Patient-level dynamics are therefore characterised by a within-host model in which the
88 values of model parameters (describing factors such as pathogen replication as well as
89 immune responses) are the same for all patients. These parameters have either
90 assumed values that have not been derived rigorously from data [29], or have been
91 obtained by fitting the model to data collected in longitudinal cohort studies from a small
92 number of patients [27,33]. Within-host parameters are, in fact, likely to vary between

93 individuals [14,39,40], for example due to differences in immune responses [41], while
94 measurement error may also lead to inaccurate parameter estimates particularly given
95 limited numbers of observations [35,42]. As we show, if patient-level data are only
96 available from a limited number of patients, then predictions of population-scale
97 epidemic dynamics may be inaccurate.

98

99 In this paper, we introduce a novel framework for transitioning from within- to between-
100 host epidemiological dynamics straightforwardly. Our method involves using a
101 compartmental model with a large number of infected compartments to predict the
102 population-scale dynamics. Compartmental models, comprising systems of ordinary
103 differential equations (ODEs), can be solved easily using standard numerical routines
104 and software packages [4,28,43,44], are straightforward to adapt to include further
105 biological detail [4,20,28], and are widely used for epidemic modelling [4,45]. We show
106 rigorously that our modelling framework is equivalent to a more complex IDE approach,
107 in the large-compartment limit of our method. Since the number of compartments is
108 simply a choice for the user to make, our easy-to-use method can generate results that
109 are as accurate as those from more complex approaches.

110

111 To demonstrate our framework, we consider modelling an outbreak due to the influenza
112 A virus. We use a previously parameterised within-host model [14] to generate a
113 synthetic dataset representative of real patient data (figure S1), incorporating variability
114 in the viral load time series between patients due to factors such as differences in
115 immune responses. Since the magnitude of this variability has been chosen to match

116 data from a previous cohort study [14,46], our dataset is comparable to obtaining data
117 from a cohort study, but with the advantage that we can test our approach using many
118 different possible cohorts of any size (from small cohort sizes up to very large cohort
119 sizes that generate idealised data). We explore the effects of both the number of
120 patients from which data are available, and the extent of measurement error in patient
121 data, on population-scale predictions. Our work therefore provides insight into the data
122 required for accurate forecasting using multi-scale epidemic models, as well as an
123 accessible modelling framework that can be used for forecasting during future
124 epidemics of a range of infectious diseases.

125

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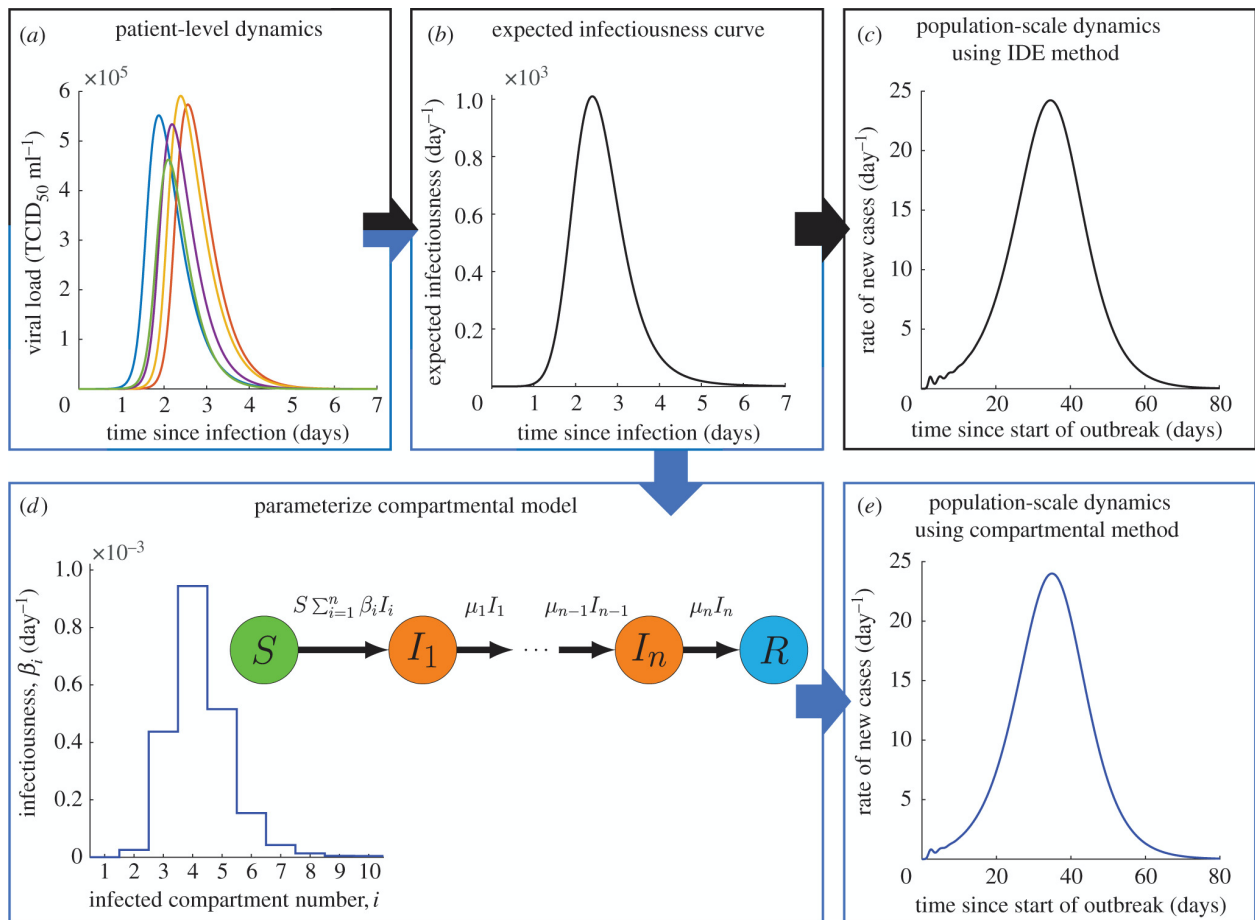
2. RESULTS

127 **Transitioning from within- to between-host influenza dynamics**

128

129 We have developed a new compartmental framework for transitioning from within- to
130 between-host epidemic dynamics. In our approach, a within-host model is fitted to data
131 from individual patients, to estimate the pathogen load of each measured patient at
132 every time since infection (figure 1a). As with other multi-scale epidemic models
133 [18,19,35], by assuming a functional relationship between pathogen load and
134 infectiousness, the expected infectiousness, $\beta(\tau)$, of any host is then estimated at each
135 time since infection, τ days (figure 1b). We will call $\beta(\tau)$ the *expected infectiousness*
136 *curve*. In previous approaches, the Kermack and McKendrick (K&M) IDE model [47]
137 (see Section S1 of the Supplementary Material) has then been used to calculate the
138 population-scale dynamics (figure 1c). We instead use the expected infectiousness

139 curve to parameterise a multi-stage compartmental model with a large number of
 140 infected compartments (figure 1d), which can also be used to predict the population-
 141 scale dynamics (figure 1e). For details on the compartmental and IDE approaches, see
 142 Methods. In the limit of infinitely many compartments in our framework, the two
 143 approaches are mathematically equivalent (we prove this rigorously in Section S2).
 144



145
 146 **Figure 1. Schematic demonstrating methods for transitioning from patient-level to population-**
 147 **scale epidemiological dynamics.** In standard approaches (a-c), measurements from within each patient
 148 (a) are used to parameterise a within-host model, giving rise to an averaged infectiousness curve (b).
 149 These patient-level dynamics can then be nested in an IDE model (e.g. the K&M model [47]) used to
 150 predict the population-scale dynamics (c). However, IDE models are challenging to solve. In contrast, in

151 our approach (*a,b,d,e*) the expected infectiousness curve is instead used to parameterise a
152 compartmental model (*d*) that can be used to predict population-scale dynamics straightforwardly (*e*).
153 Early-epidemic oscillations in panels (*c*) and (*e*) occur because the expected infectiousness of an infected
154 host is close to zero in the first day of infection, leading to delays before successive generations of newly
155 infected hosts begin to transmit the pathogen.

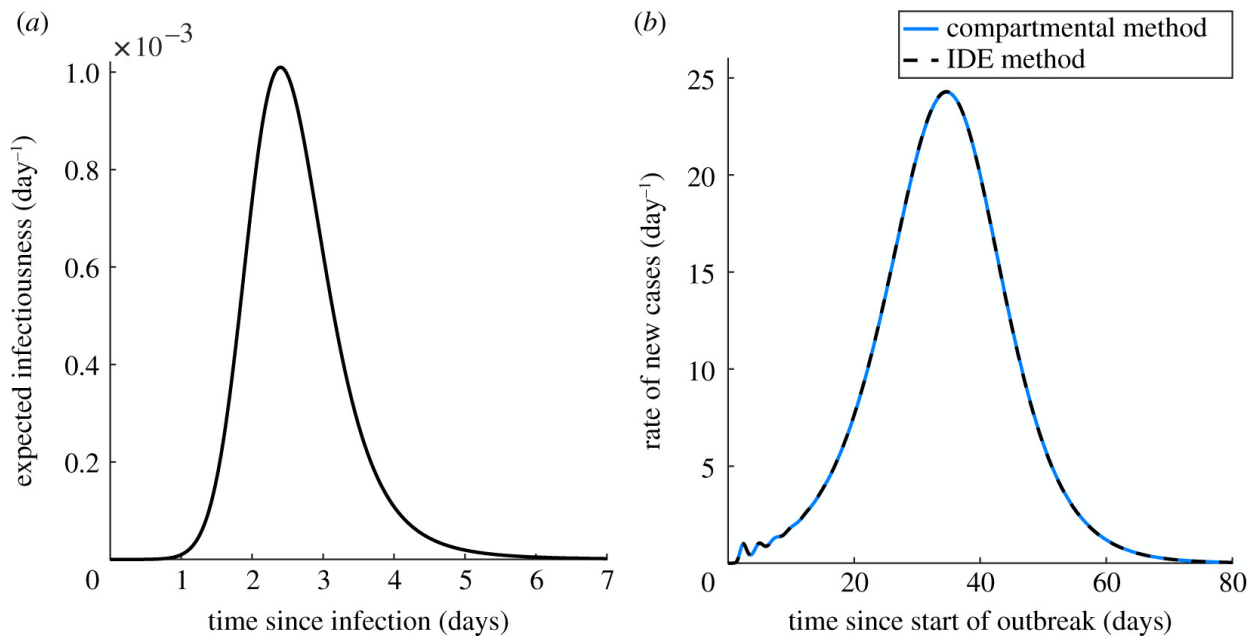
156

157 To illustrate our framework in a concrete setting, we considered the specific case of
158 influenza A infection in humans. We used the target cell-limited (TCL) within-host
159 model, which has previously been fitted to data from a cohort study of influenza
160 infection [14], to generate synthetic data from a large number of patients (see Methods).
161 The data were used to calculate the expected infectiousness curve (figure 2*a*) under the
162 assumptions of a linear relationship between viral load and infectiousness
163 [18,33,35,38,39] and a basic reproduction number (defined to be the expected number
164 of secondary cases arising from a single infected host in an otherwise entirely
165 susceptible population [4]) of 1.5 [8] (see Methods), although we consider other
166 assumptions and values of the reproduction number later (Sections S8 and S9).

167

168 Both our compartmental approach using a large number of infected compartments ($n =$
169 1000) and the previously used IDE method (i.e. the K&M model) were then used to
170 predict the population-scale outbreak dynamics, initially assuming the expected
171 infectiousness curve was known exactly (figure 2*b*). We considered a population of size
172 $N = 1000$ and assumed a single newly infected individual was introduced into an entirely
173 susceptible population. The two approaches produced almost identical results – the
174 error in the predicted population-scale dynamics when the compartmental method was

175 used, calculated as a proportional error relative to the dynamics predicted using the IDE
 176 method (see Methods), was only 0.2%. We explored how many compartments are
 177 required in our framework to ensure accurate population-scale forecasts, finding that in
 178 general, the error in predictions scales with $1/n$ as the number of compartments, n ,
 179 becomes large (Section S5). When the infectiousness curve shown in figure 2a was
 180 used to transition to population-scale dynamics, we found that $n = 24$ compartments are
 181 sufficient for an error in population-scale predictions of 10% or less (figure S2b).
 182



183
 184 **Figure 2. Transitioning from within- to between-host influenza dynamics using the compartmental**
 185 **and IDE methods.** (a) The expected infectiousness curve, $\beta(\tau)$, when the patient-level dynamics are
 186 perfectly characterised. (b) The population-scale dynamics, using our compartmental approach with $n =$
 187 1000 infected compartments (blue), and using the IDE method (black dashed), for the infectiousness
 188 curve shown in panel (a).

189

190 **The effect of limited and inaccurate patient-level data on population-scale**
191 **predictions**

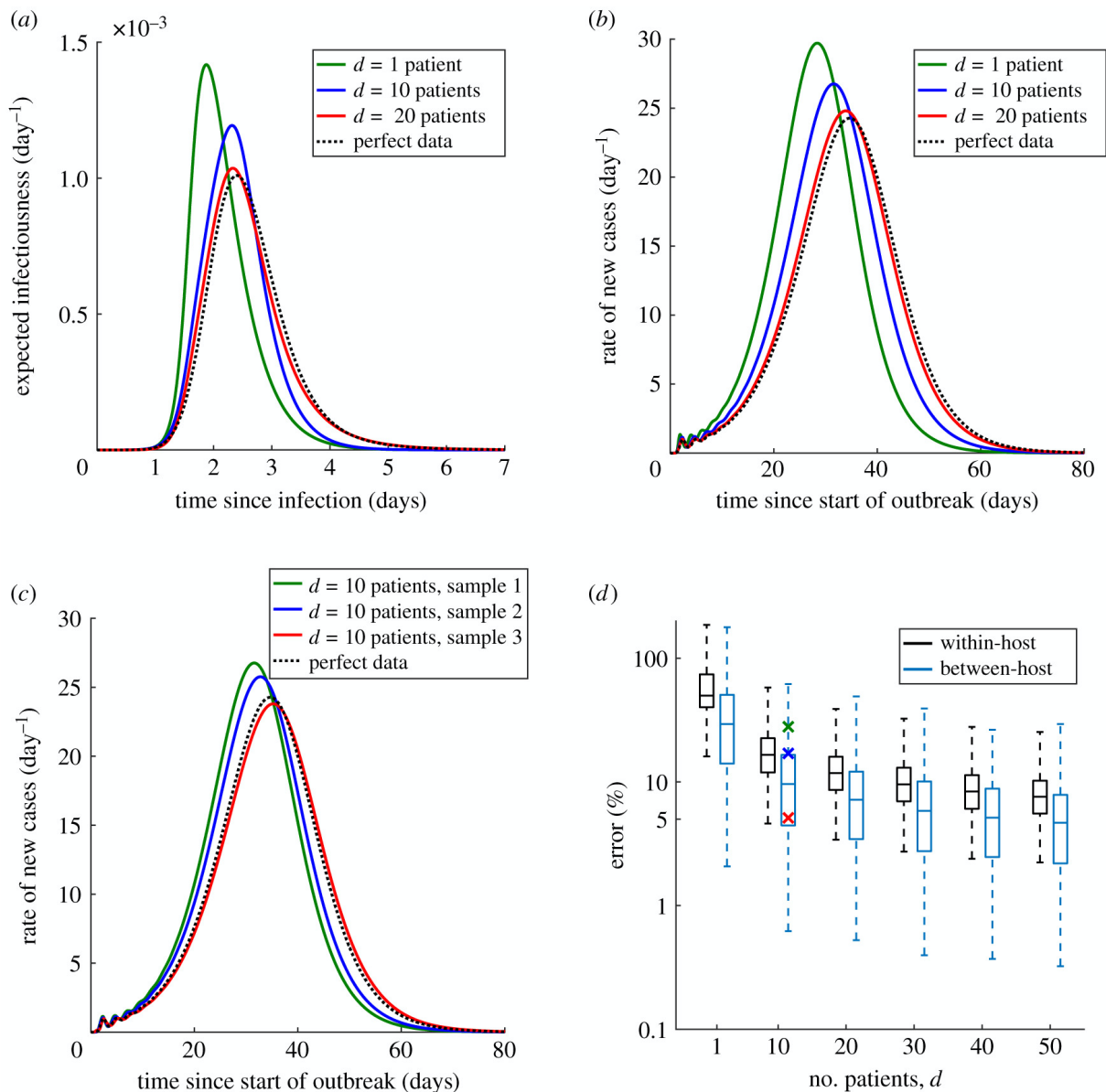
192
193 We considered the effect of two quantities on population-scale predictions: the number
194 of patients from which individual patient data are available, and the extent of
195 measurement error in patient-level data. Initially, we considered these two factors in
196 isolation, before testing their combined effects. We defined error metrics to quantify the
197 errors that arose in the patient-level dynamics (the within-host error) and in the
198 population-scale dynamics (the between-host error), using proportional errors in order to
199 enable comparison between errors at the two scales (see Methods).

200
201 *Number of patients*

202
203 In most cohort studies used to inform multi-scale models, data are only available from a
204 small number of patients [27,33] and within-host parameters may vary significantly
205 between patients [14,39,40]. To investigate the error in population-scale predictions that
206 inadequate data may generate, we supposed that data were only available from d
207 randomly chosen patients (see Methods). To isolate the effect of variability between
208 hosts rather than measurement error, the exact viral load of each patient was initially
209 assumed to be known at every time since infection. We used the available data to
210 estimate the expected infectiousness curve (figure 3a) and calculated the approximate
211 population-scale dynamics using our compartmental framework with $n = 1000$ infected
212 compartments (figure 3b). For a fixed cohort size, d , significantly different predictions of

213 population-scale dynamics are possible, depending on which patients are included in
 214 the study (figure 3c). Therefore, we calculated the distributions of errors in both the
 215 patient-level and population-scale dynamics, relative to the case in which the patient-
 216 level dynamics were perfectly characterised, over 5000 repeats for each of a range of
 217 patient cohort sizes, d (figure 3d). Equivalent results using the IDE method rather than
 218 our compartmental approach are shown in figure S3a.

219



220

221 **Figure 3. How many patients need to supply data for accurate population-scale predictions?** (a)
222 Examples of approximate expected infectiousness curves when (exact and continuous) data are available
223 from $d = 1$ (green), $d = 10$ (blue) or $d = 20$ (red) randomly chosen patients, and the expected
224 infectiousness curve when the patient-level dynamics are perfectly characterised (black dotted). (b) The
225 predicted population-scale dynamics for each infectiousness curve in panel (a), using $n = 1000$
226 compartments in our framework. (c) Three examples of possible approximate population-scale dynamics,
227 when data are available from $d = 10$ patients. (d) Box-and-whisker plots indicating the distributions of
228 within-host (black) and between-host (blue) errors for different groups of patients randomly chosen in the
229 study cohort, for a range of values of the number of patients, d . The boxes indicate the lower quartile,
230 median and upper quartile, and the maximum length of each whisker is 1.5 times the interquartile range.
231 The crosses represent the between-host errors corresponding to the curves of the same colour in panel
232 (c) (these are at values of 28%, 17% and 5% error).
233

234 As the number of patients is increased, the errors at patient-level and population-scale
235 both decrease in general (depending on precisely which patients are included in the
236 study cohort), but at a decreasing rate. The magnitude of the population-scale error is
237 generally smaller than that of the patient-level error. Therefore, limited data do not
238 necessarily preclude accurate population-scale predictions, even when there is a large
239 amount of variability between different patients. In this case – when there are exact and
240 continuous data available from each patient – a cohort size of $d = 20$ patients is
241 sufficient for the between-host error to be 10% or less on average (figure 3d). However,
242 since the errors are affected by the precise patients included in the cohort, more
243 patients are required for a greater certainty of a small between-host error. For example,
244 $d = 30$ patients are required to ensure that the upper quartile of between-host errors for
245 cohorts of that size is less than 10%.

246

247 *Extent of measurement error*

248

249 In longitudinal cohort studies, data are only collected from each patient at a limited
250 number of time points. For studies of influenza infections, data may be collected daily
251 (for example [14,39]) over the course of infection, which lasts approximately one week
252 [48]. However, there can be significant measurement error whenever the viral load is
253 recorded [35,42].

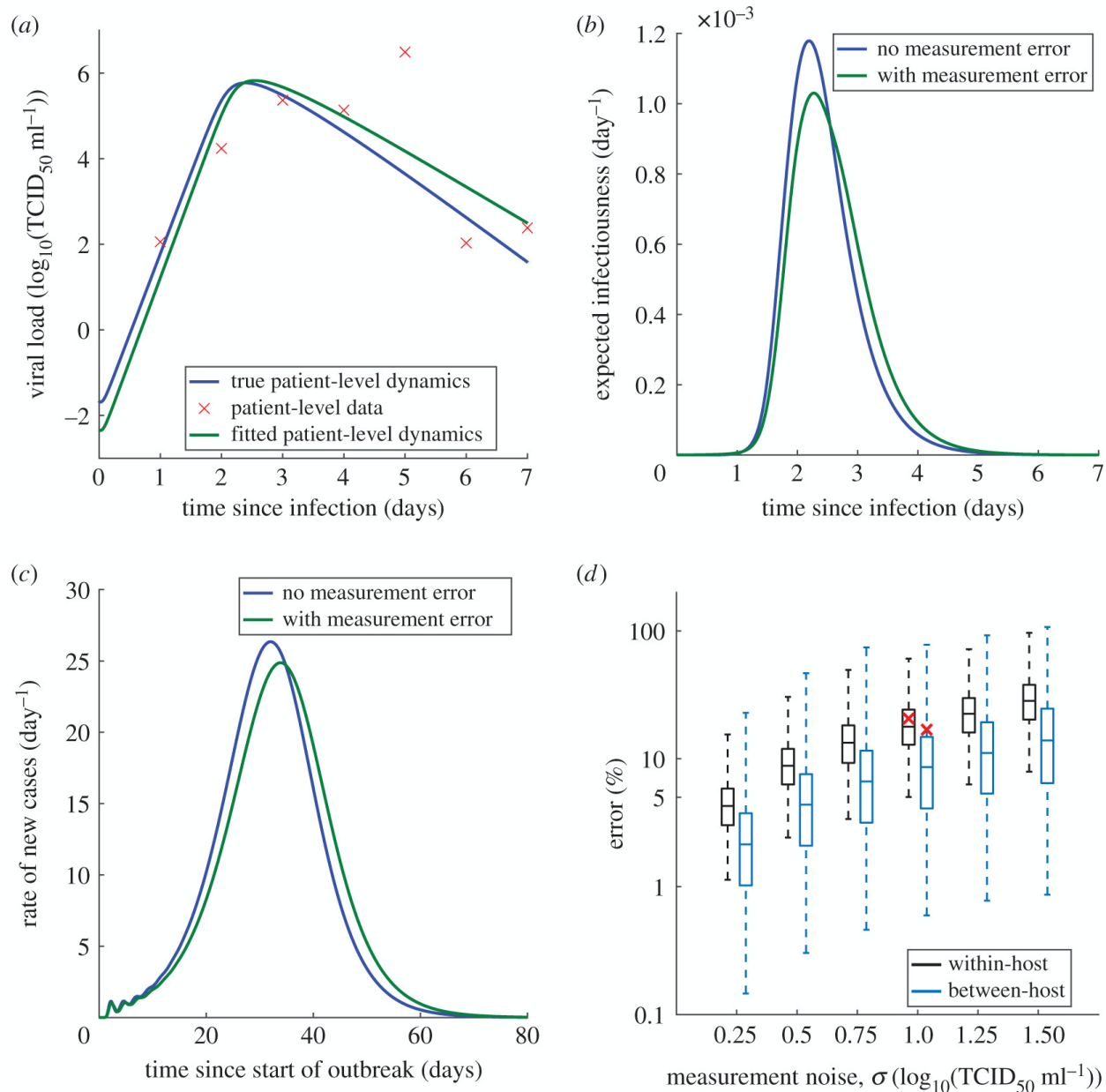
254

255 We considered viral load values recorded daily for each patient for a week after
256 infection, although we also considered the effect of more frequent observations (Section
257 S7). To incorporate measurement error in the synthetic data, a normally distributed error
258 with standard deviation σ was applied to the logarithm of each measurement. To
259 estimate the patient-level dynamics, we fitted the TCL model to the data for each patient
260 (figure 4a, see Methods for details).

261

262 To demonstrate the effect of measurement error on population-scale predictions, we
263 assumed that data were available from $d = 10$ randomly chosen patients, and compared
264 estimates of the expected infectiousness curve, first under the assumption that the viral
265 load of each host was known exactly at all times during infection, and second when
266 there was measurement error in daily recordings of the viral load (figure 4b). Our
267 compartmental framework with $n = 1000$ compartments was then used to predict the
268 population-scale dynamics in both cases (figure 4c). We calculated the within-host and

269 between-host errors that arose directly due to measurement error, by taking the “true”
 270 dynamics to be those when exact and continuous data were available from the same 10
 271 hosts. The distributions of these errors, each time calculated over 5000 repeats for a
 272 range of values of σ , are shown in figure 4d (for equivalent results obtained using the
 273 IDE method, see figure S3b).
 274



275

276 **Figure 4. The effect on population-scale predictions of measurement error in patient-level data.** (a)
277 Example of synthetic data for a single patient: the true viral load of the patient against time since infection
278 (blue), daily synthetic data with a measurement noise level of $\sigma = 1 \log_{10}(\text{TCID}_{50}/\text{ml})$ (red crosses), and
279 the viral load against time when the TCL model is fitted to the data (green). In figure S1, synthetic viral
280 load data generated using the TCL model are compared to the real data that were used to parameterise
281 the TCL model [14]. (b) Examples of expected infectiousness curves, without measurement error (blue)
282 and with measurement error (green), for $d = 10$ patients. (c) The population-scale dynamics for each
283 infectiousness curve in (b), using $n = 1000$ compartments in our framework. (d) Box-and-whisker plots
284 indicating the distributions of within-host (black) and between-host (blue) errors arising directly due to
285 measurement error, for a range of values of the extent of measurement error, σ . The red crosses
286 represent the within-host error corresponding to panel (b) and the between-host error corresponding to
287 panel (c) (these are at values of 21% and 17% error, respectively).

288
289 The errors at patient-level and at population-scale both increase with the measurement
290 noise level, σ . For values of σ of $1 \log_{10}(\text{TCID}_{50}/\text{ml})$ or higher, the mean population-scale
291 error is over 10%. In that case, when a cohort of only $d = 10$ patients is used,
292 measurement error alone is likely to prevent accurate population-scale forecasts, even if
293 there is no additional error contribution due to within-host parameter variability.

294 *Overall error*

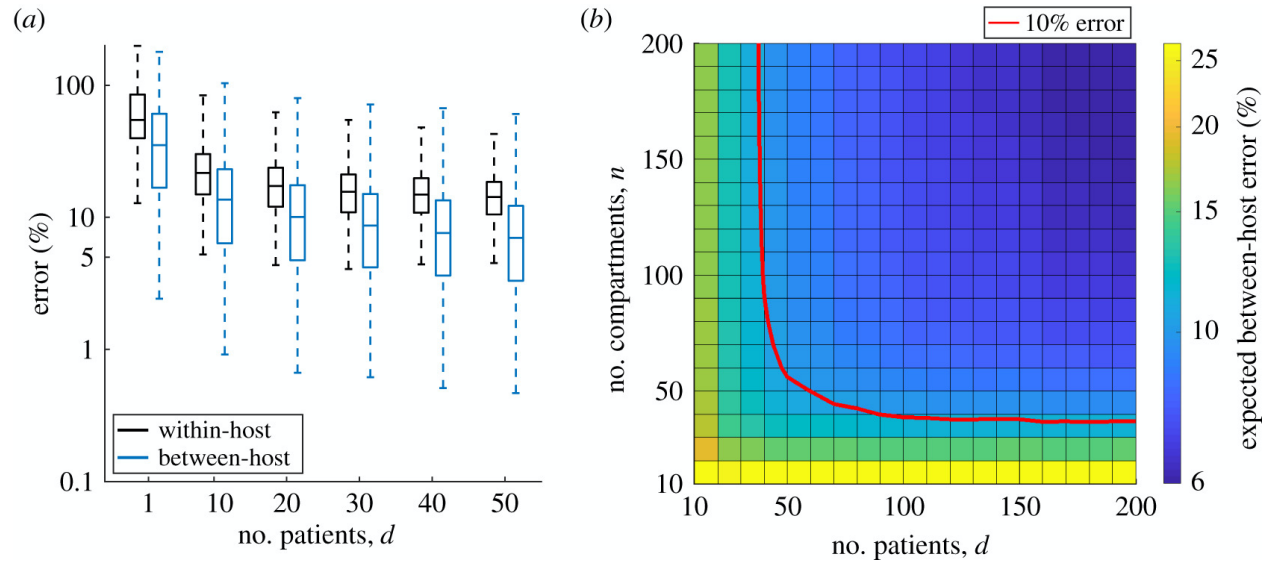
295
296
297 So far, we have described our analyses considering the separate effects of patient
298 cohort size and measurement noise on the characterisation of within-host viral load time
299 series, as well as the resulting impact on population-scale outbreak predictions.
300 However, in reality, both these sources of error would be present simultaneously. Errors

301 would also occur if a small number of compartments is used in our multi-scale modelling
302 approach, although this can be avoided by simply choosing a large number of
303 compartments in the model. Nonetheless, we also conducted an analysis in which all
304 three potential sources of error were included: (i) number of patients; (ii) measurement
305 error; (iii) number of compartments.

306

307 When we investigated the combined effect of these potential sources of error, we
308 considered a measurement noise level of $\sigma = 1 \log_{10}(\text{TCID}_{50}/\text{ml})$, since this generated
309 synthetic data comparable to those recorded in cohort studies (figure S1). Assuming
310 that data were available from d randomly sampled patients, our compartmental
311 framework with n infected compartments was used to estimate the population-scale
312 dynamics. We repeated this analysis 10,000 times each for different pairs of values of d
313 and n , each time calculating the within-host and between-host errors, relative to the
314 case in which the patient-level dynamics were perfectly characterised and the IDE
315 method was used (equivalent to using infinitely many compartments in our
316 compartmental framework). The distributions of within-host and between-host errors,
317 when a large number of compartments ($n = 1000$) is used in our framework, are plotted
318 for different numbers of patients (d) in figure 5a. Equivalent results using the IDE
319 method are shown in figure S3c. When either the compartmental or IDE approach is
320 used, data from $d = 40$ patients are required for an average between-host error of 10%
321 or below (compared to 20 patients if data are recorded exactly, i.e. with no
322 measurement error – see figure 3d and figure S3a).

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Figure 5. The effects of the number of patients and number of compartments on population-scale predictions. (a) Box-and-whisker plots indicating the distributions of within-host (black) and between-host (blue) errors for different patients chosen in the study cohort when $n = 1000$ compartments are used in our framework, assuming a measurement noise level of $\sigma = 1 \log_{10}(\text{TCID}_{50}/\text{ml})$, for a range of values of the number of patients, d . (b) The expected error in the population-scale dynamics, for different values of the number of compartments, n , and the number of patients, d . The red line indicates where the error is 10%.

The mean error in the population-scale dynamics, for different numbers of patients (d) and numbers of compartments (n), is shown in figure 5b. In the case of $d = 40$ patients, each sampled once daily, the user should choose at least $n = 60$ compartments in our approach to ensure a mean error of 10% or less. As few as 40 compartments are needed if data are available from a large number (more than 60) of patients. Nonetheless, since the number of compartments to use is simply a choice for the user – rather than requiring any more data to be collected – we suggest that any user of our framework simply chooses a very large number of compartments. We note, however,

341 that the benefit of using more compartments becomes negligible when more than $n =$
342 100 compartments are included in our approach (figure 5b).

343

344 Whereas in figure 5 we assumed data were collected once daily for a week from each
345 patient, we also conducted supplementary analyses of the between-host error when
346 data were instead collected twice per day from each patient (Section S7) – in this case,
347 data are only required from 20 patients for a mean between-host error of 10% or less
348 (figure S4b). If instead the total number of measurements that can be taken is fixed,
349 then it might be necessary to choose between sampling a large number of patients
350 infrequently, or a small number of patients frequently. We explored this in Section S7,
351 and found that sampling patients more than twice per day tended to lead to less
352 accurate population-scale predictions when the total number of measurements was
353 fixed at values below 1000 (figure S5). For realistic cohort sizes, population-scale errors
354 were similar when data were collected either once daily from $2d$ patients or twice daily
355 from d patients (for example, when $d = 20$, the respective errors are both 10%).

356

357 We examined the robustness of our results to our assumptions when transitioning from
358 within- to between-host (Section S8), finding similar results to those shown in figure 5 in
359 two alternative cases, in which the infectiousness of each patient either scales with the
360 logarithm of their viral load (figure S6c) or saturates at high viral loads (figure S6f). In
361 addition, we considered the effect of the assumed value of the basic reproduction
362 number, R_0 , on our results (Section S9), and also repeated our analyses in figure 5 for
363 different values of the measurement noise level, σ (Section S10), and for different levels

364 of variability in the within-host parameter values corresponding to different patients
365 (Section S11). When R_0 , σ , or the level of variability in within-host parameter values,
366 exceeded the values considered in figure 5, a cohort size larger than 40 hosts was
367 found to be required to ensure a mean between-host error of 10% or below (figures S7-
368 S9) – for example, data from 70 patients are required if $R_0 = 3$ (figure S7i).

369

370

3. DISCUSSION

371 In this article, we have introduced a novel, easy-to-use, compartmental framework for
372 nesting patient-level data in population-scale epidemic models. In the large-
373 compartment limit, our method is mathematically equivalent to more complicated
374 approaches that involve IDEs (Section S2). However, our method has the advantage
375 that it can be used straightforwardly, allowing it to be applied widely in future. We have
376 provided adaptable computing code alongside this article to facilitate future use of our
377 approach (see Data Accessibility).

378

379 To illustrate our method, we considered the example of influenza A infection in humans.
380 A viral dynamics model [14] was used to generate a synthetic dataset describing
381 changing viral loads in a cohort of patients, which is representative of real patient data
382 (figure S1). We showed how our compartmental framework can be used to predict the
383 population-scale epidemic dynamics, and compared our predictions to forecasts using
384 the more complicated K&M IDE model. The population-scale predictions from our
385 framework closely matched those obtained using the IDE model, provided that a

386 sufficient number of compartments was employed in our approach (figure 2b and figure
387 S2).

388

389 The amount of data used in modelling studies of within-host influenza dynamics has
390 varied widely, with some studies using data from fewer than 10 patients [14] but others
391 more than 40 patients [39]. While multi-scale models have often been parameterised
392 using either no or limited data [29], drawing robust population-scale conclusions from
393 cohort studies involving a small number of patients is likely to be challenging, since
394 patient-level dynamics display significant variability between different individuals
395 [14,39,40]. We therefore assessed the errors that arise in predicted population-scale
396 dynamics as a result of limited patient data, as well as considering measurement errors
397 that can beset parameter inference from patient-level data [42]. We first investigated
398 these effects separately (figures 3 and 4), before considering both these effects in a
399 single combined analysis (figure 5). When patient data were collected once daily, we
400 found that data from at least 40 patients were required for a mean population-scale
401 error of 10% or smaller when either our compartmental approach or the IDE method
402 was used (figure 5a and figure S3c). However, since the precise value of the
403 population-scale error depended on the exact subset of patients that was included in the
404 study, the error could be either larger or smaller than 10% even when data were
405 available from 40 patients (figure 5a). As a result, larger numbers of patients can
406 increase the confidence that the error is below a pre-specified threshold value (figure
407 5a). We considered daily measurements of pathogen load, since this frequency of data
408 acquisition is common to a number of previous longitudinal studies of influenza

409 infections (e.g. [14,39]). However, the accuracy of population-scale predictions depends
410 on the frequency with which data are collected (Section S7), so ensuring regular data
411 collection from each patient in future cohort studies is important for accurate population-
412 scale forecasting.

413

414 Our approach was motivated by earlier studies in which compartmental models with
415 multiple latent and infectious stages were employed so that the standard assumption of
416 exponentially distributed latent and infectious periods was relaxed [49–53]. The use of
417 multiple stages allows for gamma distributed latent and infectious periods (the so-called
418 “linear chain trick” [52] or “method of stages” [53]), and gamma distributions have been
419 shown to characterise epidemiological periods accurately [49,51]. However, in those
420 studies [49–53], the level of infectiousness is assumed constant throughout the
421 infectious period. We were therefore also inspired by previous research in which time-
422 dependent infectiousness was incorporated into multi-stage compartmental models, in
423 cases where the compartments correspond to clearly distinct phases of infection (for
424 example, studies of HIV [54] and Ebola [20]) or convenient time periods [55]. Our
425 approach is more similar to a method used to include experimental data in models of
426 plant disease [28], but differs from previous literature [20,28,54,55] due to our use of a
427 large number of infected compartments corresponding to different infection rates in
428 order to provide an easy method to transition from within- to between-host that is
429 accurate for any patient-level infection dynamics.

430

431 We focussed on the case study of influenza A in humans because compartmental
432 models are frequently used to model both patient-level and population-scale influenza
433 dynamics, while there has also been significant interest in developing models linking the
434 dynamics at the two scales [35]. In principle, however, our approach could be extended
435 to model outbreaks of a range of other pathogens for which patient-level dynamics are
436 well characterised. This would require careful consideration of the functional relationship
437 between pathogen load and infectiousness, since this is likely to differ between
438 pathogens [18]. In particular, the mode of transmission may be an important factor in
439 determining suitable relationships for different pathogens.

440

441 To describe individual patient-level influenza dynamics, we used the simple TCL within-
442 host model. More detailed within-host models exist, and involve features including a
443 delay before target cells begin to shed virus (an eclipse phase) [14] or explicit modelling
444 of innate and adaptive immune responses [56]. While the TCL model was sufficient to
445 demonstrate our approach here, the expected infectiousness curve in our framework
446 could be generated using a within-host model with any level of complexity. Alternatively,
447 if patient-level infection dynamics are not well characterised, then an expected
448 infectiousness curve that is estimated from transmission data [12,57], rather than within-
449 host data, could also be embedded within our framework.

450

451 In order to generate synthetic patient-level data, we assumed that two within-host
452 parameters varied between patients, using previous parameter estimates to determine
453 the level of parameter variability [14]. We incorporated measurement error by adding a

454 normally distributed random variate to daily observations of the logarithm of the viral
455 load (although we also considered other frequencies of data collection in Section S7).
456 Differences in both the extent of measurement error, and the extent of parameter
457 variability between patients, can lead to significant differences in population-scale errors
458 (figures S8 and S9). Therefore, when our modelling framework is used to determine
459 how many patients should be included in future cohort studies, careful consideration of
460 the measurement error and the variability in pathogen load time series between patients
461 is important.

462

463 The TCL model was fitted to the data from each patient using a basic least squares
464 estimation approach, since the precise method of parameter inference is not central to
465 our modelling framework. However, it would be straightforward to extend our approach
466 to consider different error structures and methods for fitting models to patient-level data.
467 In particular, a non-linear mixed effects modelling approach – amounting to a partial
468 pooling of the data between individuals – could be used. This would enable robust
469 parameter estimation in a real dataset, particularly in settings in which the numbers of
470 data points per patient are small, and both the frequency and timing of data collection
471 may vary between patients [58,59]. Going forward, we will use such a method to explore
472 further whether or not there is an optimal balance between the number of patients and
473 the frequency of measurements per patient, if total resources are limited (see Section
474 S7).

475

476 In our main analyses, we made the common assumption that the infectiousness of an
477 influenza-infected host is proportional to their viral load [18,33,35,38,39], although we
478 also obtained similar results in two alternative cases in which infectiousness either
479 scales with the logarithm of the viral load [33,38] or saturates at high pathogen loads
480 [21,55] (Section S8). However, more complex possibilities could easily be incorporated
481 into our framework. For example, future studies may also incorporate varying symptoms
482 during infection into our approach [20,23,39], in order to account for dependency of
483 transmissibility on behavioural factors in addition to pathogen load [18].

484

485 While we considered errors in population-scale predictions arising due to variability
486 between different infected patients when data are limited, our results were obtained
487 using a population-scale model in which the population was assumed to be
488 homogeneous and well-mixed. Variability between different patients was assumed to be
489 random, so that all infected hosts could effectively be assumed to follow the same
490 averaged infectiousness curve. In Section S3, we provide mathematical justification for
491 this averaging in the population-scale dynamics (see also [60]). We sought to develop
492 our framework for transitioning from within- to between-host using the simplest possible
493 population-scale model, but our compartmental approach could be generalised, for
494 example, to models incorporating age structure, spatial effects, social contact networks
495 or stochasticity [4]. In an age-structured model, different within-host parameter values
496 (or even different models) could be used to describe patient-level dynamics in the
497 different age groups, since there may be substantial differences in within-host dynamics
498 between patients of different ages [61].

499

500 In summary, we have introduced a novel compartmental framework for nesting patient
501 data in population-scale epidemiological models. We have demonstrated our easy-to-
502 use approach in the context of influenza. Not only can our modelling approach be used
503 to inform population-scale predictions with data from patients, but it can also be used to
504 design cohort studies by determining which data need to be collected. As a result, clear
505 communication between clinical epidemiologists who conduct cohort studies and
506 epidemiological modellers will allow for optimal study design. Including patient-level
507 dynamics in population-scale epidemiological models as proposed here has the
508 potential to improve epidemic forecasts; we hope that the simplicity of our approach will
509 facilitate its use for forecasting in a wide range of future outbreaks.

510

511

4. METHODS

512 **Within-host model**

513

514 The TCL model of viral dynamics, which has previously been used to model influenza
515 infections [14,15,62], is given by

$$\begin{aligned}\frac{dT}{d\tau} &= -\beta TV, \\ \frac{dI}{d\tau} &= \beta TV - \delta I, \\ \frac{dV}{d\tau} &= pI - cV,\end{aligned}\tag{4.1}$$

516 where $T(\tau)$ is the number of susceptible target cells, $I(\tau)$ is the number of infected target
517 cells, $V(\tau)$ TCID₅₀/ml is the quantity of free virus, and τ days is the time since infection.

518 The model has previously been parametrised [14] for influenza A infection in humans
 519 (Table 1).

520

521 Table 1. Estimated parameter values and initial conditions for the TCL within-host model [14].

Parameter	Definition	Value
β	Infection rate of susceptible cells by virus	$2.7 \times 10^{-5} \text{ (TCID}_{50}/\text{ml)}^{-1} \text{ day}^{-1}$
δ	Death rate of infected cells	4.0 day^{-1}
ρ	Viral shedding rate by infected cells	$1.2 \times 10^{-2} \text{ (TCID}_{50}/\text{ml)} \text{ day}^{-1}$
c	Clearance rate of free virus	3.0 day^{-1}
$T(0)$	Initial number of susceptible cells	4×10^8
$I(0)$	Initial number of infected cells	0
$V(0)$	Initial quantity of free virus	$9.3 \times 10^{-2} \text{ TCID}_{50}/\text{ml}$

522

523 We used the TCL model to generate synthetic data from different patients. To
 524 incorporate variability between patients, we assumed that the parameters δ and $V(0)$ in
 525 the TCL model vary between individuals. This represents variation in the strength of the
 526 immune response and in the initial viral load. For each patient, $\log_{10}(\delta)$ was sampled
 527 from a normal distribution with mean $0.60 \log_{10}(\text{day}^{-1})$ and standard deviation 0.25
 528 $\log_{10}(\text{day}^{-1})$, and $\log_{10}(V(0))$ was sampled from a normal distribution with mean -1.03
 529 $\log_{10}(\text{TCID}_{50}/\text{ml})$ and standard deviation $1.12 \log_{10}(\text{TCID}_{50}/\text{ml})$. These values were
 530 chosen to match variability in previous individual parameter estimates [14], while the
 531 lognormal distribution was used to guarantee positivity. All other parameters were fixed
 532 at the values given in Table 1.

533

534 We considered analyses in which viral load was assumed to be observed exactly and
535 continuously throughout infection, as well as analyses in which measurements of the
536 viral load were recorded once daily for one week after infection. In the latter case, we
537 incorporated measurement error by applying a normally distributed random variate with
538 standard deviation σ to the logarithm of each measurement. We fitted the TCL model to
539 the daily data from each patient using least squares estimation – in particular, the
540 values of the parameters δ and $V(0)$ were chosen to minimise the sum of squares
541 distance between the logarithm of the viral load in the model and in the data, while all
542 other parameter values were assumed to be known exactly and were fixed at the values
543 in Table 1. To avoid unrealistically large estimates of the initial viral load, we imposed
544 $V(0) \leq 10^3$ TCID₅₀/ml when we fitted the parameters. An example of synthetic data
545 generated for a single host, in addition to the fitted TCL model, is given in figure 4a.

546

547 **The SI_nR model**

548

549 The population-scale SI_nR model [28,63] of pathogen transmission in a population of N
550 hosts is given by

$$\begin{aligned}
\frac{dS}{dt} &= -S \sum_{j=1}^n \beta_j I_j, \\
\frac{dI_1}{dt} &= S \sum_{j=1}^n \beta_j I_j - \mu_1 I_1, \\
\frac{dI_i}{dt} &= \mu_{i-1} I_{i-1} - \mu_i I_i, \quad \text{for } i = 2, \dots, n, \\
\frac{dR}{dt} &= \mu_n I_n,
\end{aligned} \tag{4.2}$$

551 where $S(t)$ is the number of susceptible individuals, $I_i(t)$ is the number of individuals in
552 the i^{th} infected compartment, and t days is the time since the start of the outbreak.
553 Individuals in the i^{th} infected compartment infect susceptible hosts at total rate $\beta_i I_i S$ per
554 day, and progress to the next infected compartment (or recover, if $i = n$) at total rate $\mu_i I_i$
555 per day. The basic reproduction number of this model is [63]

$$R_0 = N \sum_{i=1}^n \frac{\beta_i}{\mu_i}. \quad (4.3)$$

557

558 **From within- to between-host**

559

560 We used both an existing IDE approach (steps A-C below) and a new compartmental
561 framework (steps A-B and D-E below) to transition from patient-level to population-scale
562 dynamics. The two methods are outlined below, and a schematic is shown in figure 1.

563

564 A. Fit a within-host model to longitudinally sampled data on patient-level dynamics,
565 to estimate the pathogen load of each individual patient at every time since
566 infection.

567 B. Estimate the expected infectiousness curve, $\beta(\tau)$, at each time since infection, τ
568 days, by assuming that the infectiousness of each host depends on the pathogen
569 load according to a pre-specified relationship between these quantities.

570 Then either C:

571 C. Solve the K&M IDE model, with infectiousness curve $\beta(\tau)$, to calculate the
572 population-scale dynamics (details of the K&M model are given in Section S1).

573 Or D-E:

574 D. Parameterise the SI_nR model: choose the number of infected compartments, n ,
575 where n is assumed to be large. Then find T such that $\beta(\tau)$ is zero or very small
576 for $\tau > T$ days, and choose the parameters in the SI_nR model to be

$$\begin{aligned} 577 \quad \mu_i &= \frac{n}{T}, \\ 578 \quad \beta_i &= \frac{n}{T} \int_{(i-1)T/n}^{iT/n} \beta(\tau) d\tau, \quad \text{for } i = 1, \dots, n-1, \\ 579 \quad \beta_n &= \frac{n}{T} \int_{(n-1)T/n}^{\infty} \beta(\tau) d\tau. \end{aligned} \tag{4.4}$$

580 Explanation of these parameter choices is given in Section S2.

581 E. Solve the SI_nR model numerically to approximate the population-scale dynamics.

582

583 In most of our analyses, we assumed a linear relationship between the viral load and
584 infectiousness of each influenza-infected host, although two alternative possibilities are
585 considered in Section S8. In particular, in our main analyses we assumed that

$$586 \quad \beta^{(i)}(\tau) = kV^{(i)}(\tau), \tag{4.5}$$

587 for constant k , where i represents the particular host under consideration. Therefore, the
588 expected infectiousness, $\beta(\tau)$, was given in terms of the expected viral load, $V(\tau)$, of a
589 host at time τ days since infection (calculated over a large number of realisations of the
590 within-host model), by

$$591 \quad \beta(\tau) = kV(\tau). \tag{4.6}$$

592 We fixed the constant k by assuming that the basic reproduction number,

$$593 \quad R_0 = N \int_0^{\infty} \beta(\tau) d\tau, \tag{4.7}$$

594 was known. In our main analyses, we fixed $R_0 = 1.5$, which is consistent with estimates
595 for influenza A infection [8] (different values of R_0 are considered in Section S9). The
596 expected infectiousness could therefore be calculated using the formula

$$597 \quad \beta(\tau) = \frac{R_0}{N \int_0^\infty V(x) dx} V(\tau). \quad (4.8)$$

598

599 To calculate the “true” expected infectiousness curve, $\beta(\tau)$, we computed the expected
600 viral load over 10,000 realisations of the within-host model. We also considered
601 analyses in which data were only available from a smaller number of patients, d . In such
602 cases, we simulated the within-host model d times to calculate the exact patient-level
603 dynamics corresponding to each patient, and used the data to estimate first $V(\tau)$ and
604 then $\beta(\tau)$. In analyses where we also incorporated measurement error, we used the
605 patient-level dynamics estimated by fitting the within-host model to daily observations of
606 the viral load for each patient, in order to estimate $\beta(\tau)$.

607

608 Both the compartmental and IDE methods were then used to predict the population-
609 scale dynamics. To parameterise the SI_nR model, we took $T = 7$ days, since the
610 expected infectiousness was found to be very small after a week since infection. We
611 considered a population of size $N = 1000$, and assumed that there was initially a single
612 newly infected individual, with all others susceptible. These initial conditions were
613 implemented in the SI_nR model by taking $I_1(0) = 1$ and $S(0) = 999$, with all other
614 compartments containing zero hosts initially.

615

616 **Errors at patient-level and population-scale**

617

618 We defined error metrics in order to quantify the errors that arise in the patient-level
619 dynamics and in the population-scale dynamics. These were defined as proportional
620 errors, so as to enable comparison between errors at the different scales.

621

622 First, we defined the within-host error, E_{wh} , to be the difference between the exact and
623 approximate infectiousness curves, integrated over the entire course of infection, as a
624 proportion of the area of the exact infectiousness curve. Therefore,

$$625 \quad E_{wh} = \frac{\int_0^{\infty} |\beta_{\text{approx}}(\tau) - \beta_{\text{exact}}(\tau)| d\tau}{\int_0^{\infty} \beta_{\text{exact}}(\tau) d\tau}, \quad (4.9)$$

626 where $\beta_{\text{exact}}(\tau)$ and $\beta_{\text{approx}}(\tau)$ are the exact and approximate infectiousness curves,
627 respectively.

628

629 Similarly, if $S_{\text{exact}}(t)$ and $S_{\text{approx}}(t)$ are the exact and approximate numbers of susceptible
630 individuals at time t days since the start of the epidemic, then we defined the between-
631 host error, E_{bh} , in terms of the rate of new cases per day throughout the epidemic, i.e.

$$632 \quad E_{bh} = \frac{\int_0^{\infty} |\dot{S}_{\text{approx}}(t) - \dot{S}_{\text{exact}}(t)| dt}{\int_0^{\infty} -\dot{S}_{\text{exact}}(t) dt}, \quad (4.10)$$

633 where the dot denotes differentiation with respect to time.

634

635 **AUTHOR CONTRIBUTIONS**

636 RNT conceived the research; All authors designed the study; WSH carried out the
637 research; RNT, PKM and CAY supervised the research; WSH and RNT drafted the
638 manuscript; All authors revised the manuscript and gave final approval for publication.

639

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645

646 **DATA ACCESSIBILITY**

647 All analyses were performed in MATLAB. Code is available for running the models, and
648 is available at <https://github.com/will-s-hart/WithinBetweenHostCompartmental>. Our
649 approach can also be recoded and adapted straightforwardly in other computing
650 languages.

651

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657

658

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