Assessment of disease progression in dysferlinopathy

A 1-year cohort study

Ursula Moore, MBBchir, Marni Jacobs, PhD, Meredith K. James, PT, Anna G. Mayhew, PT, PhD, Roberto Fernandez-Torron, MD, Jia Feng, MSc, Avital Cnaan, PhD, Michelle Eagle, PT, PhD, Karen Bettinson, MSc, Laura E. Rufibach, PhD, Robert Muni Lofra, PT, Andrew M. Blamire, PhD, Pierre G. Carlier, MD, PhD, Plavi Mittal, PhD, Linda Pax Lowes, PT, PhD, Lindsay Alfano, DPT, Kristy Rose, PT, PhD, Tina Duong, MPT, Katherine M. Berry, PT, Elena Montiel-Morillo, PT, Irene Pedrosa-Hernández, PT, Scott Holsten, PT, Mohammed Sanjak, PT, PhD, Ai Ashida, DPT, Chikako Sakamoto, PT, Takayuki Tateishi, PT, Hiroyuki Yajima, PT, Aurélie Canal, PT, Gwenn Ollivier, PT, Valerie Decostre, PhD, Juan Bosco Mendez, MD, Nieves Sánchez-Aguilera Praxedes, PT, Simone Thiele, PT, Catherine Siener, PT, MHS, Jeanine Shierbecker, PT, Julaine M. Florence, PT, MHS, DPT, Bruno Vandevelde, Brittney DeWolf, PT, DPT, PCS, Meghan Hutchence, BSc, PT, Richard Gee, MPT, Juliana Prügel, PT, Elke Maron, PT, Heather Hilsden, BA, Hanns Lochmüller, MD, Ulrike Grieben, MD, Simone Spuler, MD, Carolina Tesi Rocha, MD, John W. Day, MD, Kristi J. Jones, MD, PhD, Diana X. Bharucha-Goebel, MD, Emmanuelle Salort-Campana, MD, Matthew Harms, MD, Alan Pestronk, MD, Sabine Krause, MD, PhD, Olivia Schreiber-Katz, MD, Maggie C. Walter, MD, MA, Carmen Paradas, MD, PhD, Jean-Yves Hogrel, PhD, Tanya Stojkovic, MD, Shin'ichi Takeda, MD, PhD, Madoka Mori-Yoshimura, MD, PhD, Elena Bravver, MD, Susan Sparks, MD, Jordi Díaz-Manera, MD, PhD, Luca Bello, MD, PhD, Claudio Semplicini, MD, PhD, Elena Pegoraro, MD, PhD, Jerry R. Mendell, MD, Kate Bushby, MD, and Volker Straub, MD, for the Jain COS Consortium

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Abstract

Objective

To assess the ability of functional measures to detect disease progression in dysferlinopathy over 6 months and 1 year.

Methods

One hundred ninety-three patients with dysferlinopathy were recruited to the Jain Foundation's International Clinical Outcome Study for Dysferlinopathy. Baseline, 6-month, and 1-year assessments included adapted North Star Ambulatory Assessment (a-NSAA), Motor Function Measure (MFM-20), timed function tests, 6-minute walk test (6MWT), Brooke scale, Jebsen test, manual muscle testing, and hand-held dynamometry. Patients also completed the ACTIVLIM questionnaire. Change in each measure over 6 months and 1 year was calculated and compared between disease severity (ambulant [mild, moderate, or severe based on a-NSAA score] or nonambulant [unable to complete a 10-meter walk]) and clinical diagnosis.

Results

The functional a-NSAA test was the most sensitive to deterioration for ambulant patients overall. The a-NSAA score was the most sensitive test in the mild and moderate groups, while the 6MWT was most sensitive in the severe group. The 10-meter walk test was the only test showing significant change across all ambulant severity groups. In non-ambulant patients, the MFM domain 3, wrist flexion strength, and pinch grip were most sensitive. Progression rates did not differ by clinical diagnosis. Power calculations determined that 46 moderately affected patients are required to determine clinical effectiveness for a hypothetical 1-year clinical trial based on the a-NSAA as a clinical endpoint.

Conclusion

Certain functional outcome measures can detect changes over 6 months and 1 year in dysferlinopathy and potentially be useful in monitoring progression in clinical trials.

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NCT01676077.

From the John Walton Muscular Dystrophy Research Centre (U.M., M.K.J., A.G.M., R.F.-T., M.E., K.B., R.M.L., H.H., H.L., K.B., V.S.), Newcastle University and Newcastle Hospitals NHS Foundation Trust, MRC Centre for Neuromuscular Diseases, Institute of Genetic Medicine, Central Parkway, Newcastle Upon Tyne, UK; Center for Translational Science (M.J., J.F., A. Cnaan), Division of Biostatistics and Study Methodology, Cooperative International Neuromuscular Area (Grup (T.D., B.D.), and Department of Neurology (D.X.B.-G.), Children's National Health System; Pediatrics, Epidemiology and Biostatistics (M.J., A. Cnaan), George Washington University Washington, DC; Neuromuscular Area (R.F.-T.), Biodonostia Health Research Institute, Neurology Service, Donostia University Hospital, Donostia-San Sebastian, Spain; Jain Foundation (L.E.R., P.M.), Seattle, WA; Magnetic Resonance Centre (A.M.B.), Institute of Myology, Pité-Salpëtriere University Hospital, 47-83, Paris, France; Research Institute at Nationwide Children's Hospital (L.P.L., L.A., K.M.B., J.D.N.), Aneurology, Pola Alto, CA; Physical Medicine and Rehabilitation (E.M.-H., I-H.J.), Didirderi S Hospital at Westmead, University of Sydney, Australia; Lucile Salter Packard Children's Hospital at Stanford (T.D.), 24349, Neurology, Palo Alto, CA; Physical Medicine (E.M.-H.), H.-H.J., Hospital de Isanta Creu I Sanfa Creu I Stanfa Creu I Stan

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The Jain COS Consortium coinvestigators are listed in the Appendix at the end of the article.

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Correspondence Prof. Straub volker.straub@ newcastle.ac.uk

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Glossary

a-NSAA = adapted North Star Ambulatory Assessment; **COS** = Jain Clinical Outcome Study in Dysferlinopathy; **HHD** = handheld dynamometry; **LGMD2B** = limb girdle muscular dystrophy 2B; **MFM** = Motor Function Measure; **MMT** = manual muscle testing; **6MWT** = 6-minute walk test; **SRM** = standardized response mean.

Dysferlinopathy is a rare, recessively inherited muscular dystrophy caused by mutations in the $DYSF^{1-4}$ gene. Potential therapies are in development, and some of them have entered human clinical trials (NCT02579239, NCT01863004, NCT02710500). As the development of newly licensed and emerging treatments for Duchenne muscular dystrophy^{5,6} and spinal muscular atrophy type 1⁷ has shown, proving efficacy and satisfying regulators can be difficult in slowly progressive and varied conditions.⁸

Appropriate powering of a clinical trial requires understanding of disease progression and responsiveness of various outcome measures over time.⁹ Selection of outcome measures for clinical trials in dysferlinopathy is particularly challenging. Dysferlinopathy is characterized by a range of ages at onset and patterns of weakness and severity,^{4,10} variable rate of progression,⁴ modifying factors that are not clearly elucidated,^{10–13} and no appropriate measures for monitoring progression. While some measures such as the Motor Function Measure (MFM) have been shown to be reliable in dysferlinopathy,¹⁴ the variability of progression makes it difficult to demonstrate responsiveness in small cohorts.

The Jain Clinical Outcome Study in Dysferlinopathy (COS) aims to address these difficulties by characterizing the clinical, biochemical, and radiologic parameters of 209 patients with dysferlinopathy over 3 years.

This article summarizes findings from a planned interim analysis of baseline, 6-month, and year 1 data to determine whether disease progression is detectable and which measures are most able to detect this progress. We propose a power calculation using the most sensitive measures to determine a potential clinical trial cohort required to demonstrate significant functional change.

Methods

The COS study

The COS study is an international collaborative study of patients with a genetic and/or protein assay–confirmed diagnosis of dysferlinopathy.⁴ Screening, baseline, 6-month, and 1-year visits took place between October 2012 and March 2016. One hundred ninety-seven of the original 209 recruited patients completed a baseline visit. Of the original 209 patients, 7 did not meet inclusion criteria, and 5 chose not to continue past the screening visit (figure 1). Full inclusion and exclusion criteria, the study protocol, and patient demographics have been described previously.⁴

Standard protocol approvals, registrations, and patient consents

The study was initially approved by the following: NRES Committee North East—Newcastle & North Tyneside on February 2, 2012 (reference 211/NE/0360/R&D 5918). The study was also approved by ethics review boards in each country. All patients provided written informed consent. The trial was registered at ClinicalTrials.gov (NCT01676077).

Functional outcome measures

Clinical assessments were standardized through trainings at investigator meetings, and all evaluations were performed by trained clinicians. Assessments included the adapted North Star Ambulatory Assessment (a-NSAA),⁴ MFM-20, timed function tests (timed rise from floor, 10-m walk/run, 4-stair climb and descend, Timed Up and Go, and the 6-minute walk test [6MWT]), Brooke Upper Extremity Scale, Jebsen Hand Function Test, manual muscle testing (MMT), and hand-held dynamometry (HHD). No assessments have been specifically validated in a dysferlin-specific population. These assessments were chosen because of their widespread suitability for limb girdle muscular dystrophy. The MFM-20 was selected over the MFM-32 to reduce patient fatigue and duplication. All assessments were attempted for ambulant patients. Nonambulant patients did not complete the a-NSAA or ambulation-based timed function tests.

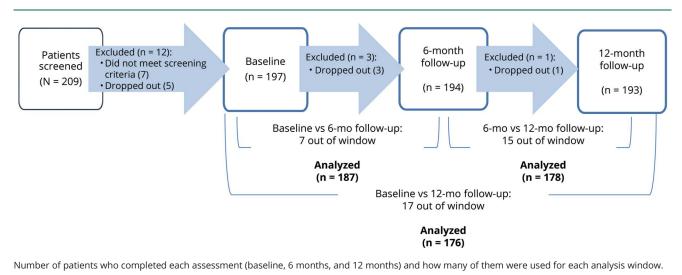
Muscle strength was assessed with an 11-point scale for MMT.⁴ The best of 3 attempts for HHD was used for analysis.

Functional ability was assessed with qualitative scales and timed tests. Total scores were calculated for the a-NSAA¹⁵ in ambulant patients, and the MFM-20 and the Brooke scale were calculated in all patients. The MFM-20 was also analyzed on the basis of its 3 components: D1 (standing and transfers), D2 (axial and proximal), and D3 (distal). Any patient with an incomplete score was excluded from analysis for that test. To account for patients who became nonambulant or unable to complete timed function tests due to disease progression, timed test values were converted to velocity measures (either meter per second or task per second), with a 0-m/s velocity assigned for the first instance a patient was unable to complete the test.

Patient perception of progression

The ACTIVLIM questionnaire, a Rasch-built patientreported outcome measure,¹⁶ was used to capture patientreported functional ability. The total score, out of a maximum of 36, was compared between visits. A higher score indicates greater functional ability.





Statistical analysis

Statistical analysis was completed with SAS 9.3 (SAS Institute Inc, Cary, NC). Mean and median change scores were calculated for each test between baseline and 6 months, 6 months and year 1, and baseline and year 1. Data from some visits were excluded if the visit occurred ±2 months outside the scheduled appointment date. A majority of functional outcome measures were not normally distributed; thus, comparisons between baseline and follow-up time points were conducted with the Wilcoxon signed rank-sum test and standardized response mean (SRM).¹⁷ A comparison was made between the performance of patients with the most common clinical diagnoses of limb girdle muscular dystrophy 2B (LGMD2B) and Miyoshi myopathy with Kruskal-Wallis test with Steel-Dwass-Critchlow-Fligner correction for multiple comparisons.¹⁸

Nonambulant patients (unable to walk 10 meters with usual orthotics and walking aids) were compared to 3 categories of ambulant patients: ambulation affected severely, moderately, or mildly. Using patients' a-NSAA scores at baseline, we defined ambulant severity categories based on a clinical impression of disease severity and binomial distribution of a-NSAA scores, with cut points of 0 to 10 (severe disease severity), 11 to 40 (moderate disease severity), and 41+ (mild disease severity). The lower cut point was based on the end of the first a-NSAA peak, and the high cut point was placed at the start of the second peak, with the wide spread in between representing a plateau between peaks. Significance was defined as $p \le 0.05$. Results are expressed as median change (range) and p value.

Power analyses to support a hypothetical future clinical trial were conducted to estimate the sample size required to detect significant statistical differences in progression over the course of 1 year with PASS 14 software. Because change scores were normally distributed, calculations used mean change scores change over time. For this analysis, the a-NSAA was used, which limited estimates to ambulant patients only. Sample sizes needed to evaluate a variety of expected drug effects were calculated for reference, including halting of disease progression (i.e., assuming mean progression estimates from the current sample among untreated and no progression among treated patients), reflective of potential protein-restoring genetic therapies. An α level of 0.05, 80% power, and a treated to-untreated ratio of 1:1 were assumed. For comparison with a widely used measure in muscular dystrophy, sample sizes were also calculated on the basis of the 6MWT.

and SDs of the functional measure that demonstrated most

Data availability statement

Deidentified cohort level data can be requested. All protocol assessments are in the public domain apart from the a-NSAA, which can be requested from the steering group (publication in draft). The statistical plan is detailed above. All data for this natural history study will be stored for at least 5 years after the end of the study. Data can be requested by all interested stakeholders for clinical research and trial planning. Data requests are reviewed by the Jain COS Steering Committee, who can be contacted via the corresponding author.

Results

Of the initial 197 eligible enrolled patients, 3 patients withdrew after baseline, and 1 withdrew after 6 months. This is summarized in figure 1. Patients included in this analysis were between 11 and 86 years of age at baseline and were at all stages of disease progression, from asymptomatic hyper-CKemia to nonambulant.

Functional outcome measures

Functional outcome measure results are summarized in tables 1 and 2, figure 2, and table e-1 (available from Dryad, doi.org/ 10.5061/dryad.tp08m60). Some of the assessments, by their

Table 1 Change in scores of functional tests from baseline to 6 months, 6 months to 1 year, and baseline to 1 year for the a-NSAA, MFM, Jebsen, Brooke, patient-reported ACTIVLIM, and 6-minute walk tests

| | Baseline | Chang | e in score, baseline–6 mc |) | Chang | e in score, 6 mo–1 y | | Chang | e in score, baseline–1 y | | |
|--|-----------------|-------|---------------------------|----------------------|-------|-----------------------|---------------------|-------|--------------------------|----------------------|------|
| Outcome measure | median score | n | Median (range) | p Value | n | Median (range) | p Value | n | Median (range) | p Value | SRM |
| Total North Star score ^a | 19 | 130 | -1.00 (-19.0 to 19.0) | <0.0001 ^a | 119 | 0.00 (-14.0 to 5.00) | 0.025 ^a | 117 | -2.00 (-16.0 to 7.00) | <0.0001 ^b | 0.61 |
| Total MFM score ^a | 45 | 161 | -1.00 (-17.0 to 7.00) | <0.0001 ^a | 148 | -0.50 (-11.0 to 11.0) | 0.0226 ^a | 149 | -1.00 (-13.0 to 6.00) | <0.0001 ^b | 0.44 |
| MFM D1 score ^a | 10 | 166 | 0.00 (-7.00 to 4.00) | 0.0007 ^a | 152 | 0.00 (-7.00 to 6.00) | 0.0033 ^a | 151 | -1.00 (-8.00 to 6.00) | <0.0001 ^b | 0.4 |
| MFM D2 score ^c | 23 | 166 | 0.00 (-12.0 to 5.00) | 0.0004 ^c | 154 | 0.00 (-6.00 to 10.0) | 0.7945 | 157 | 0.00 (-8.00 to 5.00) | 0.0171 ^b | 0.19 |
| MFM D3 score ^b | 11 | 174 | 0.00 (-3.00 to 2.00) | 0.0655 | 158 | 0.00 (-3.00 to 3.00) | 0.1431 | 160 | 0.00 (-6.00 to 2.00) | 0.0011 ^b | 0.26 |
| Brooke Upper Extremity Scale ^b | 1 | 181 | 0.00 (-2.00 to 5.00) | 0.2025 | 164 | 0.00 (-5.00 to 2.00) | 0.477 | 164 | 0.00 (-2.00 to 2.00) | 0.0334 ^b | 0.18 |
| ACTIVLIM total score ^a | 26 | 180 | 0.00 (-9.00 to 8.00) | 0.0042 ^a | 165 | 0.00 (-9.00 to 10.0) | 0.0026 ^a | 164 | -1.00 (-8.00 to 6.00) | <0.0001 ^b | 0.46 |
| Jebsen writing time taken, s | 12.7 | 180 | 0.10 (-17.3 to 14.9) | 0.5794 | 165 | 0.10 (-11.3 to 30.5) | 0.56 | 162 | -0.15 (-14.4 to 28.1) | 0.7109 | 0.11 |
| 6-min walk distance, meters/min ^c | 331 | 124 | -4.00 (-95.0 to 123) | 0.2041 | 112 | -7.00 (-110 to 51.0) | 0.0003 ^c | 113 | -13.0 (-123 to 142) | <0.0001 ^b | 0.37 |

Abbreviations: a-NSAA = adapted North Star Ambulatory Assessment; MFM = Motor Function Measure; SLM = standardized response mean. *p* Values were calculated with the Wilcoxon signed-rank method. The SRM is a measure of effect size reported for change over 1 year. ^a Consistent significant change over 6 months.

^b Change over 1 year.

^c Change over one 6-month period but not both 6-month periods.

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Table 2 Change scores for MMT with a possible range of scores of 0 to 10 and HHM measured with a dynamometer

| | Median score | Chan | ge in score, baseline–6 n | no | Chan | ge in score, 6 mo–1 year | | Chan | ge in score, baseline–1 y | | |
|---|--------------------------------|------|---------------------------|---------------------|------|--------------------------|----------------------|------|---------------------------|----------------------|------|
| | at baseline, median (range) | n | Median (range) | p Value | n | Median (range) | p Value | n | Median (range) | p Value | SRM |
| MMT shoulder flexion | 7 (1–10) | 184 | 0.00 (-3.00 to 3.00) | 0.9973 | 169 | 0.00 (-3.00 to 3.00) | 0.0577 | 166 | 0.00 (-4.00 to 3.00) | 0.1979 | 0.20 |
| MMT shoulder abduction ^a | 7 (1–10) | 184 | 0.00 (-6.00 to 5.00) | 0.2465 | 169 | 0.00 (-4.00 to 4.00) | 0.1172 | 166 | 0.00 (-4.00 to 3.00) | 0.011 ^a | 0.11 |
| HHM elbow flexion biceps, lb ^b | 16.60 (1.00–65.00) | 177 | -0.70 (-24.5 to 10.80) | 0.0012 ^b | 159 | -0.20 (-28.7 to 16.90) | 0.9424 | 156 | -0.60 (-30.6 to 15.10) | 0.0026 ^a | 0.23 |
| MMT elbow flexion biceps ^b | 8 (1–10) | 184 | 0.00 (-4.00 to 4.00) | 0.0322 ^b | 169 | 0.00 (-4.00 to 4.00) | 0.1412 | 166 | 0.00 (-4.00 to 4.00) | 0.0016 ^a | 0.13 |
| HHM elbow flexion brachioradialis, lb | 16.40 (0.00–56.60) | 175 | -0.30 (-29.1 to 17.80) | 0.1451 | 159 | -0.10 (-20.4 to 14.40) | 0.5082 | 156 | -0.45 (-32.9 to 15.20) | 0.0616 | 0.07 |
| MMT elbow flexion brachioradialis | 8 (1–10) | 184 | 0.00 (-5.00 to 3.00) | 0.7979 | 169 | 0.00 (-4.00 to 4.00) | 0.1109 | 166 | 0.00 (-5.00 to 3.00) | 0.1388 | 0.24 |
| MMT elbow extension ^a | 8.00 (2-10) | 176 | 0.00 (-5.00 to 4.00) | 0.1263 | 161 | 0.00 (-3.00 to 5.00) | 0.9587 | 160 | 0.00 (-5.00 to 4.00) | 0.0472 ^a | 0.03 |
| MMT wrist flexion ^a | 9 (2–10) | 184 | 0.00 (-7.00 to 5.00) | 0.0647 | 169 | 0.00 (-8.00 to 6.00) | 0.3377 | 166 | 0.00 (-4.00 to 3.00) | 0.0024 ^a | 0.15 |
| HHM wrist flexion, lb | 15.80 (0.00–46.10) | 176 | 0.00 (-18.1 to 15.10) | 0.721 | 161 | -0.20 (-22.4 to 17.30) | 0.4597 | 158 | -0.15 (-28.3 to 15.30) | 0.8098 | 0.04 |
| HHM wrist extension, lb | 20.40 (3.10-50.30) | 178 | -0.15 (-16.3 to 20.30) | 0.5521 | 163 | 0.00 (-25.0 to 15.40) | 0.9028 | 160 | -0.55 (-26.9 to 17.60) | 0.162 | 0.23 |
| MMT wrist extension | 9 (3–10) | 184 | 0.00 (-3.00 to 4.00) | 0.4453 | 169 | 0.00 (-3.00 to 4.00) | 0.5773 | 166 | 0.00 (-3.00 to 3.00) | 0.3273 | 0.16 |
| HHM grip, Ib ^b | 60.00 (0.00-260.0) | 172 | -1.00 (-173 to 103.1) | 0.1287 | 160 | -1.00 (-83.0 to 94.40) | 0.0058 ^b | 152 | -2.00 (-186 to 94.90) | 0.0004 ^a | 0.01 |
| HHM pinch grip, lb ^b | 24.80 (0.00-97.00) | 171 | 0.00 (-42.0 to 39.00) | 0.2784 | 158 | -1.00 (-35.0 to 22.00) | <0.0001 ^b | 150 | -1.00 (-34.0 to 40.00) | 0.0019 ^a | 0.20 |
| MMT hip flexion ^a | 6 (0–10) | 176 | 0.00 (-7.00 to 4.00) | 0.3953 | 160 | 0.00 (-4.00 to 7.00) | 0.1663 | 160 | 0.00 (-5.00 to 4.00) | 0.0093 ^a | 0.18 |
| MMT hip extension | 3 (0–10) | 173 | 0.00 (-4.00 to 5.00) | 0.5764 | 156 | 0.00 (-4.00 to 5.00) | 0.9785 | 155 | 0.00 (-7.00 to 4.00) | 0.6247 | 0.04 |
| MMT hip abduction | 6 (1–10) | 175 | 0.00 (-5.00 to 7.00) | 0.2452 | 156 | 0.00 (-5.00 to 5.00) | 0.5303 | 156 | 0.00 (-7.00 to 7.00) | 0.5382 | 0.26 |
| HHM hip abduction, lb | 20.50 (2.10–71.80) | 161 | 0.10 (-37.4 to 34.30) | 0.4108 | 148 | 0.10 (-33.3 to 22.70) | 0.6222 | 149 | 0.10 (-36.5 to 26.10) | 0.49 | 0.21 |
| MMT hip adduction ^b | 2 (0–10) | 175 | 0.00 (-5.00 to 5.00) | 0.0732 | 156 | 0.00 (-4.00 to 4.00) | 0.0052 ^b | 156 | 0.00 (-5.00 to 5.00) | 0.0009 ^a | 0.16 |
| HHM hip adduction, lb ^a | 11.65 (0.00–49.60) | 147 | -0.10 (-25.2 to 21.30) | 0.5118 | 133 | -0.40 (-22.4 to 15.30) | 0.0838 | 138 | -1.15 (-28.4 to 17.20) | 0.0168a | |
| MMT knee flexion ^b | 4 (0–10) | 173 | 0.00 (-5.00 to 4.00) | 0.0256 ^b | 155 | 0.00 (-3.00 to 4.00) | 0.6457 | 155 | 0.00 (-5.00 to 3.00) | 0.014 ^a | 0.26 |
| HHM knee flexion, lb ^b | 9.60 (0.00-65.70) | 141 | -0.70 (-19.0 to 15.90) | 0.007 ^b | 127 | -0.40 (-34.1 to 14.80) | 0.1201 | 127 | -1.20 (-44.5 to 20.00) | <0.0001 ^a | 0.20 |
| MMT knee extension | 3 (0–10) | 182 | 0.00 (-3.00 to 4.00) | 0.5574 | 167 | 0.00 (-4.00 to 5.00) | 0.2943 | 164 | 0.00 (-5.00 to 5.00) | 0.5064 | 0.12 |
| HHM knee extension, lb ^b | 11.80 (0.00–95.30) | 154 | -0.55 (-35.2 to 22.60) | 0.0007 ^b | 139 | -0.20 (-39.5 to 41.10) | 0.3253 | 138 | -0.80 (-37.7 to 19.70) | 0.0052 ^a | 0.05 |
| MMT ankle dorsiflexion ^a | 4 (0–10) | 180 | 0.00 (-7.00 to 6.00) | 0.1944 | 165 | 0.00 (-6.00 to 7.00) | 0.1841 | 164 | 0.00 (-5.00 to 6.00) | 0.0295 ^a | 0.24 |
| HHM ankle dorsiflexion, lb ^b | 11.30 (0.00–55.90) | 140 | -0.75 (-29.5 to 22.70) | 0.0018 ^b | 118 | 0.00 (-28.8 to 30.00) | 0.7194 | 118 | -0.50 (-40.5 to 25.90) | 0.0537 | 0.23 |

| | Median score | Change | ge in score, baseline–6 mo | 0 | Chan | Change in score, 6 mo-1 year | | Chang | Change in score, baseline-1 y | | |
|--|--------------------------------|--------|----------------------------|---------------------|------|------------------------------|---------------------|-------|-------------------------------|---------------------|------|
| | at baseline, median (range) | ۶ | Median (range) | <i>p</i> Value | 5 | Median (range) | <i>p</i> Value | Ē | Median (range) | <i>p</i> Value | SRM |
| MMT ankle plantarflexion (knee straight) ^a | 3 (0–10) | 157 | 0.00 (-10.0 to 7.00) | 0.273 | 139 | 0.00 (-6.00 to 8.00) | 0.212 | 137 | 0.00 (-10.0 to 8.00) | 0.0193 ^a | 0.23 |
| HHM ankle plantar flexors (knee straight). Ib 12.60 (0.00–81.00) 140 | 12.60 (0.00-81.00) | 140 | -0.40 (-34.2 to 23.70) | 0.21 | 127 | -0.50 (-28.6 to 32.90) | 0.229 | 123 | -0.70 (-33.9 to 30.20) | 0.1945 | 0.23 |
| MMT ankle plantarflexion (knee flexed) ^a | 3 (0–10) | 173 | 0.00 (-6.00 to 7.00) | 0.2528 | 154 | 0.00 (-5.00 to 6.00) | 0.242 | 153 | 0.00 (-6.00 to 5.00) | 0.0174 ^a | 0.04 |
| HHM ankle plantar flexors (knee flexed), lb ^b | 12.20 (0.00–57.90) 144 | 144 | -0.75 (-25.6 to 29.00) | 0.0138 ^b | 127 | -0.30 (-34.5 to 22.10) | 0.2234 | 124 | 124 –1.15 (–39.6 to 22.60) | 0.002 ^a | 0.17 |
| MMT ankle inversion | 7 (0–10) | 180 | 0.00 (-7.00 to 6.00) | 0.6828 | 165 | 0.00 (-5.00 to 4.00) | 0.8154 | 164 | 0.00 (-4.00 to 5.00) | 0.6789 | 0.16 |
| MMT ankle eversion ^b | 5 (0-10) | 179 | 0.00 (-6.00 to 4.00) | 0.8232 | 164 | 0.00 (-5.00 to 4.00) | 0.0003 ^b | 164 | 164 0.00 (-6.00 to 3.00) | 0.013 ^a | 0.21 |

over both 6-month periods. The SRM, a measure of effect size, is reported for change over 1 year. MMT and HHM assessments all showed low responsiveness to change.

^o Change over 1 year.
^o Change over a single 6-month period but not both 6-month periods.

nature, are suitable only for ambulant patients. In addition, some patients did not complete all assessments because of equipment or individual patient factors such as recent injury and were excluded from analysis of that assessment. The number of patients who completed each assessment is listed.

The a-NSAA, MFM-20 (total and subdomain D1), the Timed Up and Go, and timed 10-meter walk showed significant change in median score over both 6-month periods and 1 year (table 1 and figure 2); no MMT or HHD values changed consistently over both 6-month windows (table 2).

The MFM D2, some MMT and HHD, velocity to rise from floor, climb 4 stairs, descend 4 stairs, and the 6MWT showed significant change over only one 6-month window and over 1 year (tables 1 and 2 and figure 2).

The MFM D3, some MMT and HHD, and the Brooke scale showed change in median score at 1 year but did not change over either 6-month windows (tables 1 and 2).

The timed test demonstrating the greatest change in velocity was the 10-meter walk, changing -0.1 m/s (range -1.24 to 0.55 m/s) from a baseline median velocity of 1.11 m/s (range 0.00-4.55 m/s). The velocity of all timed tests at year 1 slowed by almost 10% of baseline velocity for that task (figure 2).

The Jebsen Hand Function Test and remaining MMT and HHD tests showed no change over 1 year.

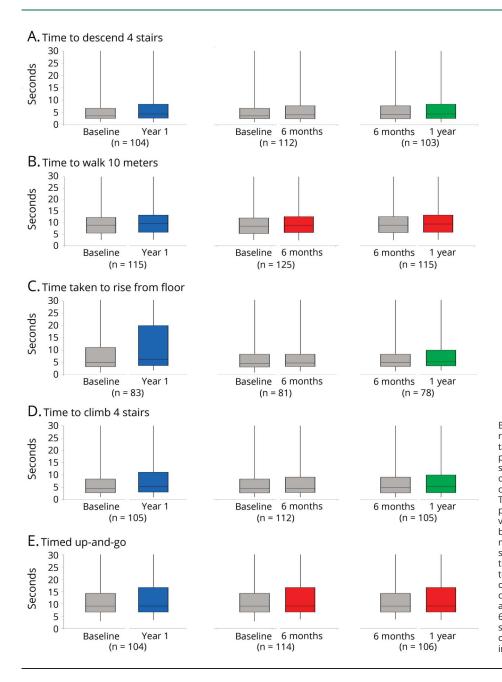
The tests most responsive to change for the whole cohort, as calculated by the SRM, were NSAA (SRM 0.61), 10-meter walk/run velocity (0.56), ACTIVLIM (0.46), rise from floor velocity (0.44), and MFM total score (0.44).

When assessed by severity subgroup, for ambulant patients, the a-NSAA was the most sensitive test in the mildly (SRM 0.44) and moderately (SRM 0.89) affected groups, while the 6MWT was most sensitive in the severely affected ambulant group (SRM 0.9). The 10-meter walk was the only test showing significant change in all ambulant subgroups. For nonambulant patients, the most sensitive test was the MMT for wrist flexion (SRM 0.69), while the most sensitive combined measure was the distal domain of the MFM (SRM 0.46). Table e-1 (available from Dryad doi.org/10.5061/dryad.tp08m60) shows SRM values for all assessments by severity subgroup. Change in functional test scores (MFM and a-NSAA) did not differ between the 2 most common clinical diagnoses of LGMD2B and Miyoshi myopathy (table 3).

Patient perception of progression

Responses recorded by the ACTIVLIM questionnaire supported the deterioration in function, with the ACTIVLIM score decreasing statistically significantly over both 6 and 12 months (table 1).





Boxplots showing range, interquartile range (IQR), and median values of time taken to perform timed tests. Paired comparisons of calculated velocity (1/time in seconds) were used to determine significant vs insignificant change in order to include those unable to complete the test. These are displayed as time taken to complete each task (rather than velocity) to aid visual interpretation. Comparisons of baseline and year 1, baseline and 6 months, and 6 months and 1 year are shown. The numbers differ depending on the number of patients completing each test at both visits. Blue shows significant change (p < 0.05) over 1 year; red shows if change is seen over both 6-month periods; and green shows change in only the green 6-month window. Some patients were very slow, and those taking >25 seconds to complete a test are not displayed but are included in median and IQR calculations.

Estimated sample size for trials

Estimated clinical trial sample size is highly dependent on the expected treatment effect; the smaller the effect size, the larger the sample required (table 4). For example, with the inclusion of ambulant patients of all disease severities, a total of 328 patients (164 in both the treatment and placebo groups) would be needed to detect a 50% reduction in disease progression (i.e., half the deterioration seen in the current sample) over 1 year compared to 90 patients (45 in each group) needed to detect halting in progression. The a-NSAA was chosen for its capability to detect change as demonstrated by the highest overall SRM in the current analysis. Limiting a clinical trial to a subset of patients in the moderate severity

group at baseline, who showed the greatest changes in a-NSAA scores over 1 year, would reduce the total sample size to 176 for a 50% reduction in progression and 46 to detect a halt in progression. A treatment that is anticipated to improve muscle function, producing a greater effect size than simply halting progression, would require even fewer patients.

The 6MWT would require 216 ambulant patients to detect a halting of progression. However, as reflected by the high SRM in the severe-ambulant population, if trials were limited to the severe-ambulant subgroup of patients, 42 patients would be needed to detect a halting in progression with the 6MWT.

Table 3 Scores in LGMD2B and Miyoshi myopathy

| | | | | | | | | | Kruskal- Wallis test of difference between |
|--|-----|------------------------|---------|------|-----|------------------------|------------|------|--|
| | LGN | 1D2B | | | Miy | oshi myopathy | | | groups |
| Outcome | n | Median (range) | p Value | SRM | n | Median (range) | p Value | SRM | p Value |
| Total North Star score | 67 | -2.00 (-14.0 to 5.00) | <0.0001 | 0.69 | 35 | -2.00 (-16.0 to 7.00) | 0.0015 | 0.50 | 0.8663 |
| Total MFM Score | 84 | -1.00 (-13.0 to 6.00) | <0.0001 | 0.46 | 48 | -1.00 (-10.0 to 6.00) | 0.0131 | 0.38 | 0.87 |
| MFM D1 score | 85 | -1.00 (-6.00 to 6.00) | 0.0006 | 0.36 | 49 | 0.00 (-8.00 to 3.00) | 0.0042 | 0.42 | 0.8898 |
| MFM D2 score | 91 | 0.00 (-8.00 to 4.00) | 0.0201 | 0.25 | 49 | 0.00 (-4.00 to 5.00) | 0.4588 | 0.08 | 0.8873 |
| MFM D3 score | 92 | 0.00 (-3.00 to 2.00) | 0.0026 | 0.32 | 51 | 0.00 (-6.00 to 2.00) | 0.2350 | 0.19 | 0.6595 |
| MMT shoulder abduction | 97 | 0.00 (-4.00 to 3.00) | 0.0530 | 0.21 | 51 | 0.00 (-2.00 to 2.00) | 0.2754 | 0.18 | 0.9883 |
| MMT shoulder flexion | 97 | 0.00 (-3.00 to 2.00) | 0.2873 | 0.12 | 51 | 0.00 (-2.00 to 3.00) | 0.8111 | 0.02 | 0.8958 |
| MMT elbow flexion biceps | 97 | 0.00 (-4.00 to 4.00) | 0.1284 | 0.16 | 51 | 0.00 (-3.00 to 3.00) | 0.0505 | 0.25 | 0.9713 |
| MMT elbow flexion brachioradialis | 97 | 0.00 (-5.00 to 3.00) | 0.3926 | 0.11 | 51 | 0.00 (-3.00 to 3.00) | 0.7598 | 0.03 | 0.9249 |
| MMT wrist extension | 97 | 0.00 (-3.00 to 3.00) | 0.9115 | 0.01 | 51 | 0.00 (-2.00 to 2.00) | 0.6209 | 0.09 | 0.8412 |
| MMT wrist flexion | 97 | 0.00 (-4.00 to 3.00) | 0.0750 | 0.19 | 51 | 0.00 (-4.00 to 3.00) | 0.0444 | 0.28 | 0.9326 |
| MMT knee extension | 96 | 0.00 (-3.00 to 4.00) | 0.9472 | 0.04 | 51 | 0.00 (-3.00 to 5.00) | 0.5997 | 0.06 | 0.9994 |
| MMT ankle dorsiflexion | 96 | 0.00 (-5.00 to 4.00) | 0.0536 | 0.20 | 51 | 0.00 (-3.00 to 6.00) | 0.1403 | 0.14 | 0.9966 |
| MMT ankle inversion | 96 | 0.00 (-4.00 to 4.00) | 0.8853 | 0.01 | 51 | 0.00 (-4.00 to 5.00) | 0.7866 | 0.04 | 0.9927 |
| MMT ankle eversion | 96 | 0.00 (-5.00 to 3.00) | 0.0701 | 0.22 | 51 | 0.00 (-6.00 to 3.00) | 0.1469 | 0.22 | 0.9941 |
| MMT ankle plantarflexion (knee straight) ^a | 78 | 0.00 (-5.00 to 8.00) | 0.6160 | 0.02 | 44 | 0.00 (-10.0 to 7.00) | 0.0062 | 0.36 | 0.0455 |
| MMT ankle plantarflexion (knee flexed) | 86 | 0.00 (-6.00 to 5.00) | 0.2716 | 0.11 | 50 | 0.00 (-2.00 to 5.00) | 0.0876 | 0.17 | 0.6968 |
| MMT hip extension | 88 | 0.00 (-5.00 to 4.00) | 0.2450 | 0.12 | 50 | 0.00 (-7.00 to 2.00) | 0.4626 | 0.37 | 0.423 |
| MMT knee flexion | 89 | 0.00 (-5.00 to 3.00) | 0.3173 | 0.14 | 50 | 0.00 (-3.00 to 2.00) | 0.0147 | 0.20 | 0.3917 |
| MMT hip Abduction | 89 | 0.00 (-7.00 to 7.00) | 0.2638 | 0.12 | 50 | 0.00 (-4.00 to 6.00) | 0.5599 | 0.10 | 0.4836 |
| MMT hip Adduction | 89 | 0.00 (-5.00 to 3.00) | 0.0025 | 0.32 | 50 | 0.00 (-5.00 to 5.00) | 0.6168 | 0.06 | 0.2017 |
| MMT hip flexion | 92 | 0.00 (-5.00 to 2.00) | 0.1230 | 0.19 | 51 | 0.00 (-3.00 to 4.00) | 0.1110 | 0.20 | 0.8431 |
| MMT elbow extension | 92 | 0.00 (-5.00 to 4.00) | 0.0653 | 0.20 | 51 | 0.00 (-3.00 to 4.00) | 0.8632 | 0.02 | 0.6868 |
| HHM elbow flexion biceps, lb | 91 | -0.30 (-17.6 to 11.00) | 0.0904 | 0.24 | 48 | -0.85 (-23.6 to 15.10) | 0.0498 | 0.23 | 0.7967 |
| HHM elbow flexion brachioradialis, lb | 91 | -0.60 (-20.5 to 8.20) | 0.2003 | 0.19 | 48 | 0.20 (-21.0 to 15.20) | 0.5768 | 0.09 | 0.9248 |
| HHM wrist extension, lb ^a | 93 | -1.40 (-17.4 to 14.60) | 0.0173 | 0.21 | 50 | 1.40 (–15.5 to 17.60) | 0.0798 | 0.15 | 0.0187 |
| HHM wrist flexion, lb | 93 | -0.10 (-16.6 to 15.30) | 0.8544 | 0.03 | 48 | 0.60 (-10.4 to 7.60) | 0.5540 | 0.05 | 0.8952 |
| HHM grip, lb | 93 | -1.00 (-186 to 87.10) | 0.0085 | 0.27 | 42 | -2.00 (-85.0 to 62.90) | 0.1015 | 0.25 | 0.9728 |
| HHM pinch grip, lb | 93 | -1.00 (-31.0 to 40.00) | 0.0175 | 0.19 | 42 | 0.00 (-27.0 to 16.00) | 0.2026 | 0.28 | 0.9623 |
| HHM knee extension, lb | 81 | -0.10 (-30.3 to 15.50) | 0.2259 | 0.17 | 42 | -1.10 (-36.6 to 19.70) | 0.0215 | 0.24 | 0.5599 |
| HHM knee flexion, lb | 73 | -1.00 (-11.6 to 16.60) | 0.0276 | 0.14 | 39 | -1.00 (-13.5 to 20.00) | 0.0011 | 0.21 | 0.6533 |
| HHM hip abduction, lb | 86 | -0.10 (-14.7 to 19.20) | 0.8929 | 0.08 | 48 | 0.90 (-36.5 to 26.10) | 0.0948 | 0.11 | 0.3955 |

Continued

| | LGN | 1D2B | | | Miy | oshi myopathy | | | Kruskal- Wallis test of difference between groups |
|--|-----|------------------------|---------|------|-----|------------------------|------------|------|--|
| Outcome | n | Median (range) | p Value | SRM | n | Median (range) | p Value | SRM | p Value |
| HHM hip adduction, lb | 82 | -1.00 (-20.8 to 13.40) | 0.1396 | 0.16 | 41 | -1.10 (-15.5 to 17.20) | 0.2555 | 0.07 | 0.9905 |
| HHM ankle dorsiflexion, lb | 68 | -0.65 (-18.5 to 17.00) | 0.0833 | 0.15 | 35 | 0.00 (-22.2 to 25.90) | 0.8751 | 0.01 | 0.5334 |
| HHM ankle plantar flexors (knee flexed), lb | 71 | -1.00 (-23.0 to 20.90) | 0.0331 | 0.24 | 37 | -1.20 (-10.7 to 16.10) | 0.1155 | 0.11 | 0.9482 |
| HHM ankle plantar flexors (knee straight), lb | 71 | -1.00 (-27.0 to 17.30) | 0.1283 | 0.18 | 36 | -0.45 (-30.6 to 14.40) | 0.9234 | 0.08 | 0.69 |
| Brooke Upper Extremity Scale | 97 | 0.00 (-1.00 to 2.00) | 0.0301 | 0.21 | 51 | 0.00 (-1.00 to 1.00) | 0.3877 | 0.17 | 0.9764 |
| ACTIVLIM total score | 95 | -1.00 (-8.00 to 4.00) | <0.0001 | 0.51 | 51 | -1.00 (-8.00 to 6.00) | 0.0520 | 0.28 | 0.8076 |
| Jebsen writing time taken, s | 95 | -0.20 (-14.4 to 28.10) | 0.6479 | 0.15 | 49 | 0.50 (-10.5 to 6.40) | 0.5693 | 0.04 | 0.9795 |
| 6-min walk test total distance, meters/min | 65 | -8.00 (-123 to 71.00) | 0.0137 | 0.33 | 34 | -15.50 (-108 to 142.0) | 0.0607 | 0.24 | 0.9881 |

Abbreviations: HHM = hand-held myometry; LGMD2B = limb girdle muscular dystrophy type 2B; MFM = Motor Function Measure; MMT = manual muscle testing; SRM = standardized response mean.

This table shows the change scores over 1 year for patients with a diagnosis of LGMD2B or Miyoshi myopathy. Change scores were compared between groups with the Kruskal-Wallis test and Steel-Dwass-Critchlow-Fligner control for multiple comparisons. The SRM is a measure of effect size demonstrated by each test, shown for each diagnostic group. Overall functional scores of the MFM and adapted North Star Ambulatory Assessment did not differ between diagnostic subgroups.

^a Only 2 muscle groups (ankle plantar flexion and wrist extension) showed a difference between these 2 groups.

Discussion

This analysis supports the use of specified physical and patientreported outcome measures in the assessment of dysferlinopathy by demonstrating clinically detectable deterioration in muscle strength and function and self-reported deterioration over 6 months and 1 year in a large cohort of patients.

Defining a set of disease specific tests is useful in coordinating further research and patient follow-up and developing clinical trial outcome measures. The most sensitive assessments for monitoring progression in dysferlinopathy that were identified in this study are the a-NSAA, MFM-20, timed 10-meter walk, and Timed Up and Go; they demonstrated consistent change over both 6 months and 1 year with a relatively high SRM. Other tests such as the Brooke and some MMT assessments showed change over 1 year but not 6 months, suggesting that they are less sensitive scales. This was also true for D3 (distal function) of the MFM, consistent with other published crosssectional data on its poor suitability for this population.¹⁴ Some tests such as the 6MWT and MFM D2 domain changed over one 6-month window but not the other. This poor repeatability of demonstrable change makes these measures less reliable in assessing the dysferlinopathy patient cohort as a whole.

The 6MWT showed the highest SRM for the severely affected ambulant patients, perhaps capturing a large deterioration in the walking pace of some of these weak patients as they lose ambulation. However, this is unlikely to be used in clinical trials, particularly because this subgroup of patients is small (37 severe-ambulant patients in this study), thus reducing the number of patients available to appropriately power a study.

Trials in rare diseases are often forced to rely on small cohort sizes. Here, we show that while variation in physical progression can be high in dysferlinopathy, sensitive outcome measures can be identified (a-NSAA, 10-meter walk, etc.), particularly in ambulant patients. We confirm that a moderate disease severity group based on the a-NSAA could be identified, making it possible to power a clinical trial with a reasonably sized sample. The actual number needed would depend on the expected treatment effect over 1 year. Numbers needed for a trial could be further reduced by using this longitudinal natural history sample as a control group for future clinical studies.¹⁹ In addition, longer evaluation times and identification of more sensitive scoring systems could further reduce the number of participants needed for effective clinical trials. Future work aiming to assess the cohort by modeling methods will look at progression based on a variety of disease and environmental factors to attempt to determine whether any subgroups exist with differing rates of progression. Any particularly fast-progressing group would be very good candidates for potential trials in terms of both patient benefit and reduced sample size requirement.

Increasing the sensitivity and utility of functional scoring systems across the disease spectrum requires the creation of

Table 4 Sample size estimates for a placebo-controlled clinical trial

| | Target treatment eff | ect, n | | | |
|---|------------------------------|------------------------------|------------------------|--------------------|--------------------|
| Target study population | 50% Reduction in progression | 75% Reduction in progression | Halting of progression | 20% Improvement | 50% Improvement |
| All ambulant patients | 328 | 148 | 90 | 62 | 42 |
| Moderate patients only (baseline a-NSAA score 11–40) | 176 | 80 | 46 | 34 | 22 |

Abbreviation: a-NSAA = adapted North Star Ambulatory Assessment.

Sample size estimates for a placebo-controlled clinical trial are based on projected mean change in a-NSAA score for variable treatment effects over 1 year. The numbers represent the total sample size based on equally sized treatment and control groups. Calculations are based on an observed decrease in NSAA score in untreated patients of 2.4 overall and 3.8 among moderate patients, the expected control effect. Estimated change in a-NSAA score among treatment groups over 1 year is therefore as follows: 50% reduction, decrease of 1.2 (all) and 1.9 (moderate) in a-NSAA score; 75% reduction, decrease of 0.6 (all) and 0.9 (moderate) in a-NSAA score; halting of progression, no change in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score; and 50% improvement, increase of 1.2 (all) and 1.9 (moderate) in a-NSAA score; balting of progression, no change in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score; halting of 1.2 (all) and 1.9 (moderate) in a-NSAA score; balting of progression, no change in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score; halting of progression, no change in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score; 20% improvement, increase of 0.5 (all) and 0.7 (moderate) in a-NSAA score.

a disease-specific scale. We are working to streamline the a-NSAA and MFM-20 to remove redundant and less sensitive items and to combine elements of each scale using a Rasch analysis²⁰ to create a linearizable, dysferlin-specific score. This modified scale, called the North Star Assessment for Dysfer-linopathy, will be validated with the 2- and 3-year COS visits and will be assessed with Rasch analysis to determine whether the new scale is capable of measuring change over a wider proportion of the patient population over time.

Characterizing which individual muscles demonstrate most rapid deterioration in dysferlinopathy is useful for both monitoring progression and targeting potential therapeutics, some of which involve local intramuscular injections.²¹ These results over 1 year show some sensitivity to change of a few distal muscle groups in nonambulant patients, in whom functional scores are less useful. However, in ambulant

Table 5Intraclass correlation coefficient between
screening and baseline assessments for
functional tests carried out at a maximum of 90
days apart

| Test | Intraclass correlation | n | Multiple |
|-------------------------|------------------------|-----|----------|
| NSAA | 0.99 | 132 | 119 |
| MFM total score | 0.99 | 170 | 150 |
| Time to stand | 0.85 | 91 | 77 |
| Time to walk/run 10 m | 0.99 | 127 | 114 |
| Timed Up and Go | 0.84 | 112 | 104 |
| Time to climb 4 steps | 0.93 | 109 | 101 |
| Time to descend 4 steps | 0.53 | 108 | 101 |
| 6-min walk distance | 0.97 | 129 | 121 |
| ACTIVLIM | 0.97 | 174 | 167 |
| | | | |

Abbreviations: MFM = Motor Function Measure; NSAA = North Star Ambulatory Assessment.

A high intraclass correlation demonstrates consistency between screening and baseline assessments. patients, there is a low responsiveness and high variability of MMT and HHD results because measures of specific muscle change. Other measures of strength, longer time scales, and MRI may be needed to monitor effects of local treatments.

In all tests, a small number of individuals of all ages and disease severities improved their time or qualitative functional score over 1 year. While in some this may reflect a true improvement in function, we would not expect an improvement in underlying pathology. Therefore, it is likely that some of this improvement can be attributed to potentially confounding factors. With functional scales, improvement may be due to a learning effect²²; that is, patients become more practiced at certain tests, although we attempted to mitigate this by having all patients perform a full assessment at screening. Some intertester variation in scoring is also possible because, although all were trained by the same provider, different physiotherapists conducted consecutive patient visits at some sites. However, we anticipate this effect to be small because the functional assessments used have previously demonstrated high inter-tester reliability in other diseases, ^{23–25} and intraclass correlation comparison of baseline and screening results (table 5) shows consistency of scoring, particularly for the a-NSAA (intraclass correlation 0.99), on which much of this analysis is based. Finally, this may simply reflect environmental, physiologic, or individual factors such as motivation or fatigue, which can cause natural fluctuations in performance on functional tests²² on the background of more slowly progressing pathology. If trends in scores are maintained in future visits, suggesting a true subgroup of static or even functionally improving patients, this will be revisited to assess for potential disease modifiers.

The ACTIVLIM is a rating scale with documented high sensitivity in many neuromuscular conditions.^{16,26} We have shown that this holds true for the dysferlinopathy population, who show changes in score across all levels of function. A very small number of patients reported an improvement in the ability to perform tasks of daily living. These patients had undertaken lifestyle changes, including weight loss, dietary change, and participation in physiotherapy. Future work will explore and capture patient access to physiotherapy,

participation in exercise, and the use of aids and adaptations in the home and workplace. These findings suggest that the ACTIVLIM is a reliable patient-reported outcome measure in dysferlinopathy that is a valuable complementary tool for clinical trials at any stage of the disease.

There are several limitations of this study. The first is the relatively short period of time over which change was examined; this will be addressed in future analysis of this 3-year study. Second, the use of ordinal, rather than continuous, functional scales presented challenges in data analysis, necessitating the use of medians, making intuitive interpretation more difficult. However, these scales also have benefits because a single point change on a functional change necessarily represents a clinically, rather than simply statistically, noticeable change. Future work will explore the potential to linearize these Rasch-based measures, allowing us to keep the benefits of these scales while providing the opportunity to subject them to more rigorous parametric statistics.

Although function is measured, the importance attributed by patients to the loss of ability to perform 1 particular action over another is not measured here. To clarify the clinical significance of the statistically significant changes demonstrated here, further correlations are needed between these functional and patient-reported measures, as well as consideration of patient valuation of the importance of the loss of individual functions.

Our current assessment of arm function is limited in the scales currently used in this study. Further work is required to develop or determine suitable measures for upper limb function.

This study demonstrates significant population-level change over 6 months and 1 year in several functional and patientreported measures. We anticipate that the planned combination of the most sensitive elements of the a-NSAA and MFM to create a dysferlin-specific score will produce an important measure that will become the basis for assessment of progression in dysferlinopathy regardless of ambulatory status. This measure combined with MRI data and further patient subgroup analysis will add to the pathophysiologic understanding and trial readiness of the dysferlinopathy population and may further reduce the cohort size required to power potential clinical trials.

Author contributions

Ursula Moore: analysis or interpretation of the data, statistical analysis, drafting or revising the manuscript for intellectual content. Marni Jacobs: statistical analysis. Meredith James, analysis or interpretation of the data, drafting or revising the manuscript for intellectual content. Anna Mayhew and Roberto Fernandez Torron: analysis or interpretation of the data, drafting or revising the manuscript for intellectual content. Jia Feng: analysis or interpretation of the data, statistical analysis. Avital Cnaan: design or conceptualization of the study, drafting or revising the manuscript for intellectual content. Karen Bettinson: major role in the acquisition of data, drafting or revising the manuscript for intellectual content. Laura Rufibach: design or conceptualization of the study, drafting or revising the manuscript for intellectual content. Robert Muni Lofra, Andrew M. Blamire, and Pierre G. Carlier: design or conceptualization of the study, major role in the acquisition of data, drafting or revising the manuscript for intellectual content. Plavi Mittal: design or conceptualization of the study, drafting or revising the manuscript for intellectual content. Michelle Eagle: design or conceptualization of the study, major role in the acquisition of data, drafting or revising the manuscript for intellectual content. Linda Lowes, Lindsay Alfano, Kristy Rose, Tina Duong, Katherine Berry, Elena Montiel-Morillo, Irene Pedrosa-Hernández, Scott E. Holsten, Mohammed Sanjak, Ai Ashida, Chikako Sakamoto, Tatyuki Tateishi, Hiroyuki Yajima, Aurélie Canal, Gwenn Olliver, Valerie Decostre, Bosco Mendez, Nieves Sánchez-Aguilera Práxedes, Simone Thiele, Catherine Siener, Jeannine Sheierbecker, Julaine Florence, Bruno Vandevelde, Brittney DeWolf, Meghan Harman, Richard Gee, Juliane Prugel, Elke Maron, and Heather Hilsden: major role in the acquisition of data, drafting or revising the manuscript for intellectual content. Hanns Lochmüller: design or conceptualization of the study, drafting or revising the manuscript for intellectual content. Ulrike Grieben, Simone Spuler, Carolina Tesi Rocha, John W. Day, Kristi J. Jones, Diana X. Bharucha-Goebel, Emmanuelle Salort-Campana, Matthew Harms, Alan Pestronk, Sabine Krause, Olivia Schreiber-Katz, Maggie C. Walter, Carmen Paradas, Jean-Yves Hogrel, Tanya Stojkovic, Shin'ich Takeda, Madoka Mori-Yoshimura, Elena Bravver, Susan Sparks, Jordi Díaz-Manera, Luca Bello, Claudio Semplicini, Elena Pegoraro, Jerry R. Mendell, and Kate Bushby: design or conceptualization of the study, major role in the acquisition of data, drafting or revising the manuscript for intellectual content. Volker Straub: design or conceptualization of the study, major role in the acquisition of data, drafting or revising the manuscript for intellectual content, overall responsibility for the content of this paper. For Study Group: the Jain Consortium: All members of the Jain Consortium are listed in the Appendix. All members of the study group played a role in acquisition of data.

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about the foundation. If you are a patient with dysferlinopathy, please consider enrolling in their interactive dysferlinopathy registry, which seeks to build a strong, engaged, and supportive community (patients@jain-foundation.org).

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Appendix 1 Coinvestigators: The Jain COS Consortium

Coinvestigator Affiliation and role in study Adrienne Arrieta, MS Children's National Medical Center, Washington, DC; data management and training **Esther Hwang** Jain Foundation, Seattle, WA; recruitment, development of assessment form Elaine Lee, PhD Jain Foundation, Seattle, WA; recruitment, development of assessment forms Isabel Illa, MD Hospital de la Santa Creu i Sant Pau/CIBERER, Barcelona, Spain; site investigator Eduard Gallardo, MD Hospital de la Santa Creu i Sant Pau/CIBERER, Barcelona, Spain; site investigator Izaskun Belmonte Jimeno, PT Servei de Medicina F ísica i Rehabilitació, Hospital de la Santa Creu i Sant Pau, Barcelona, Spain; site investigator Jaume Llauger Rossello Radiology Department, Universitat Autònoma de Barcelona, Hospital de la Santa Creu i Sant Pau, Barcelona, Spain; radiologist **Bruce Harwick** Department of Radiology, CMC Mercy Charlotte, Carolinas Healthcare System Neurosciences Institute, Charlotte, NC; NMR technologist Jackie Sykes, RN, BSN Carolinas HealthCare System, Charlotte, NC; study coordinator Brent Yetter, MS Nationwide Children's Hospital, Columbus, OH; study coordinator Mark Smith, MS, DABMP, Department of Radiology, Nationwide Children's Hospital, Columbus, OH; site investigator RT (MR) Bernard Lapeyssonie, PT Neuromuscular and ALS Center, La Timone Hospital, Aix-Marseille Université, Marseille, France; site investigator David Bendahan, PhD Centre de Résonance, MagnétiqueBiologique et Médicale, UMR CNRS 7339, Marseille, France; Aix-Marseille Université, Marseille, France; site investigator Yann Le Fur, PhD Aix-Marseille Université, Marseille, France; site investigator Attarian Shahram, MD, PhD Neuromuscular and ALS Center, La Timone Hospital, Aix-Marseille Université, Marseille, France; site investigator **Testot-Ferry Albane, CRA** Neuromuscular and ALS Center, La Timone Hospital, Aix-Marseille Université, Marseille, France; study coordinator Eva M. Coppenrath, MD Department of Clinical Radiology, Ludwig-Maximilians-University Munich, Germany; site investigator Elizabeth Harris, MD John Walton Muscular Dystrophy Research Centre, Newcastle Upon Tyne, UK; clinical investigator Michela Guglieri, MD John Walton Muscular Dystrophy Research Centre, Newcastle Upon Tyne, UK; clinical investigator Teresinha Evangelista, MD The John Walton Muscular Dystrophy Research Centre, Newcastle Upon Tyne, UK; clinical investigator Alex Murphy, MD The John Walton Muscular Dystrophy Research Centre, Newcastle Upon Tyne, UK; clinical investigator The John Walton Muscular Dystrophy Research Centre, Newcastle Upon Tyne, UK; clinical investigator **Dionne Moat** Tim Hodgson, MClinRES Magnetic Resonance Centre, Institute for Cellular Medicine, Newcastle University, Newcastle Upon Tyne, UK; site investigator Dorothy Wallace, BSc Magnetic Resonance Centre, Institute for Cellular Medicine, Newcastle University, Newcastle Upon Tyne, UK; site investigator Louise Ward, DCR Magnetic Resonance Centre, Institute for Cellular Medicine, Newcastle University, Newcastle Upon Tyne, UK; site investigator **Debra Galley** Magnetic Resonance Centre, Newcastle University, Newcastle Upon Tyne, UK; radiology assistant Chiara Calore University of Padova, Padova, Italy, site investigator Roberto Stramare, MD Radiology Unit, Department of Medicine, University of Padova, Padova, Italy; site investigator

| Appendix 1 (continued) | |
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| Coinvestigator | Affiliation and role in study |
| Alessandro Rampado, MRT | Radiology Unit, Department of Medicine, University of Padova, Padova, Italy; site investigator |
| Teresa Gidaro | Institut de Myologie, Paris, France; site investigator |
| Suna Turk, MSc | AIM & CEA NMR Laboratory, Institute of Myology, Pitié-Salpêtrière University Hospital, 47-83, Paris, France; site investigator |
| Laurent Servais | Institut de Myologie, Paris, France; site investigator |
| Cyrille Theis | Institut de Myologie, Paris, France; site investigator |
| Oumar Diabaté | Institut de Myologie, Paris, France; study coordinator |
| Linda Schimmoeller | Washington University, St. Louis, MO; study coordinator |
| Glenn Foster, RTR (MR) | Center for Clinical Imaging Research CCIR, Washington University, St. Louis, MO; site investigator |
| Pilar Carbonell, MD | Hospital U. Virgen del Rocío/Instituto de Biomedicina de Sevilla, Seville, Spain; site investigator |
| Macarena Cabrera, MD | Hospital U. Virgen del Rocío/Instituto de Biomedicina de Sevilla, Seville, Spain; site investigator |
| Yolanda Morgado, MD | Hospital U. Virgen de Valme/Instituto de Biomedicina de Sevilla, Seville, Spain; site investigator |
| Susana Rico Gala, MD | Department of Radiology, Hospital U. Virgen de Valme, Seville, Spain; site investigator |
| Jennifer Perez | Stanford University School of Medicine, Stanford, CA; study coordinator |
| Anne Marie Sawyer, FSMRT | Lucas Centre for Imaging, Stanford University School of Medicine, Stanford, CA; site investigator |
| Nigel F. Clarke, MD | Institute for Neuroscience and Muscle Research, Sydney, Australia; site investigator |
| Sarah Sandaradura, MD | Institute for Neuroscience and Muscle Research, Sydney, Australia; site investigator |
| Roula Ghaoui, MD | Institute for Neuroscience and Muscle Research, Sydney, Australia; site investigator |
| Kayla Cornett, Ex Phys | Institute for Neuroscience and Muscle Research, Sydney, Australia; site investigator |
| Claire Miller, PT | Institute for Neuroscience and Muscle Research, Sydney, Australia; site investigator |
| Sheryl Foster, MHlthSc | Department of Radiology, Westmead Hospital; Faculty of Health Sciences, University of Sydney, Australia; site investigator |
| Anthony Peduto, MBBS | Department of Radiology, Westmead Hospital; Faculty of Health Sciences, University of Sydney, Australia; site investigator |
| Noriko Sato, MD, PhD | Department of Radiology, National Center Hospital, National Center of Neurology and Psychiatry, Tokyo, Japan; site investigator |
| Takeshi Tamaru, MRT | Department of Radiology, National Center Hospital, National Center of Neurology and Psychiatry, Tokyo, Japan; site investigator |
| Yoko Kobayashi, MD | National Center of Neurology and Psychiatry, Tokyo, Japan; site investigator |
| Ai Ashida, PT | National Center of Neurology and Psychiatry, Tokyo, Japan; site investigator |
| Takahiro Nakayama, MD, PhD | Yokohama Rosai Hospital, Yokohama, Japan; study advisor |
| Kazuhiko Segawa, MD, PhD | National Center of Neurology and Psychiatry, Tokyo, Japan; site investigator |
| Sachiko Ohtaguro | National Center of Neurology and Psychiatry, Tokyo, Japan; study assistant |
| Harumasa Nakamura, MD | National Center of Neurology and Psychiatry, Tokyo, Japan; study advisor |
| Maki Ohhata | National Center of Neurology and Psychiatry, Tokyo, Japan; study coordinator |
| En Kimura, MD, PhD | National Center of Neurology and Psychiatry, Tokyo, Japan; study advisor |
| Makiko Endo | National Center of Neurology and Psychiatry, Tokyo, Japan; study coordinator |
| Nora Brody, PT, DPT | Children's National Health System, Washington, DC: site investigator |
| Meganne E. Leach, MSN, APRN | Children's National Health System, Washington, DC: site investigator |
| Allyn Toles | Children's National Health System, Washington, DC: study coordinator |
| Stanley T. Fricke, PhD | Department of Diagnostic Imaging and Radiology, Children's National Health System, Washington, DC; site investigator |
| Hansel J. Otero, MD | Department of Diagnostic Imaging and Radiology, Children's National Health System, Washington DC; site investigator |
| | |

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