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for non-communicable disease surveillance**

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1 **Advances in spatio-temporal models for non-communicable** 2 **disease surveillance**

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12

13 **Abstract**

14 Surveillance systems are commonly used to provide early warning detection or to
15 assess an impact of an intervention/policy. Traditionally, the methodological and
16 conceptual frameworks for surveillance have been designed for infectious diseases,
17 but the rising burden of non-communicable diseases (NCDs) worldwide suggests a
18 pressing need for surveillance strategies to detect unusual patterns in the data and to
19 help unveil important risk factors in this setting. Surveillance methods need to be able
20 to detect meaningful departures from expectation and exploit dependencies within
21 such data to produce unbiased estimates of risk as well as future forecasts. This has
22 led to the increasing development of a range of space-time methods specifically
23 designed for NCD surveillance.

24 We present an overview of recent advances in spatio-temporal disease surveillance
25 for NCDs using hierarchically specified models. This provides a coherent framework
26 for modelling complex data structures, dealing with data sparsity, exploiting
27 dependencies between data sources and propagating the inherent uncertainties
28 present in both the data and the modelling process. We then focus on three commonly
29 used models within the Bayesian Hierarchical Model (BHM) framework and through a
30 simulation study we compare their performance.

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3 31 We also discuss some challenges faced by researchers when dealing with NCD
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5 32 surveillance, including how to account for false detection and the modifiable areal unit
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7 33 problem. Finally, we consider how to use and interpret the complex models, how
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9 34 model selection may vary depending on the intended user group and how best to
10
11 35 communicate results to stakeholders and the general public.

12
13 36 **Keywords:** surveillance, non-communicable diseases, Bayesian hierarchical models,
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15 37 spatio-temporal modelling.
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20 21 39 **Key messages**

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24 40 - There is increasing recognition of the importance of surveillance for NCDs.
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26 41 - Spatio-temporal variation in health outcomes and lifestyle and environmental
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28 42 exposures needs to be explicitly modelled in order to reduce bias and
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30 43 uncertainty.
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32 44 - Hierarchical modelling provides a coherent framework within which spatio-
33
34 45 temporal dependencies can be explicitly modelled with integration of the
35
36 46 uncertainties associated both with the data and the modelling process.
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38 47 - In a simulation study, we found that mixture models designed for detection
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40 48 perform better than standard disease mapping models. However, attention
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42 49 should be paid to the choice of threshold as this impacts the results. It is
43
44 50 recommended that a simulation study based on the characteristics of the data
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46 51 in hand is run each time for the selection of suitable threshold values.
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48 52 - Current research challenges in this area include: the use of data from multiple
49
50 53 sources at different spatial and temporal scales and with different sources of
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52 54 bias and uncertainty; computationally intense processes; and control for false
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54 55 positive findings.
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57 **1. The importance of non-communicable diseases**

58 According to the World Health Organization surveillance is the “ongoing systematic
59 data collection, analysis, interpretation and dissemination of information in order for
60 action to be taken” [1]. National public health agencies, such as the US Centers for
61 Disease Control and Prevention (CDC) and Public Health England (PHE) routinely
62 carry out surveillance data analysis to provide early warnings of unexplained changes
63 in incidence patterns of diseases as well as to aid policy formation and resource
64 allocation [2]. Specific examples include the international influenza monitoring system
65 which started in 1948 and is now distributed in 82 countries [3], the HIV and AIDS
66 Reporting System (HARS) used by PHE [4] and the National HIV Surveillance System
67 used by CDC [4].

68 To date, the majority of methods and models commonly used in public health
69 surveillance are designed for monitoring cases of infectious diseases [6]. Due to the
70 rising burden of non-communicable diseases (NCDs) worldwide, there is a pressing
71 need to implement surveillance strategies to detect trends, highlight unusual changes
72 and consequently assist in outlining emerging NCD risk factors. NCD surveillance
73 shares many objectives with infectious disease surveillance, including generating
74 information to guide public health action, detecting the health impact of environmental
75 exposures, or of environmentally driven disease vectors; however, it also presents
76 some different methodological challenges [7, 8].

77 Health data contain both a time and a space component. Surveillance methods must
78 be able to capture spatial and temporal patterns in both lifestyle/environmental
79 exposures and health outcomes. Here, we present an overview of the approaches
80 developed for spatio-temporal disease surveillance of NCDs. We focus on model-

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3 81 based methods and among these on Bayesian hierarchical models (BHMs), which can
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5 82 naturally accommodate complex data structures, as well as propagate uncertainty due
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7 83 to the data themselves and the modelling process.
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10 84 In this Section, we first discuss how data availability is one of the key challenges in
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12 85 surveillance studies, before giving a generic overview of test-based approaches for
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14 86 NCD surveillance. We then focus on BHMs (Section 2) and describe disease mapping
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16 87 and mixture-based models for anomaly detection. In Section 3, we introduce the
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18 88 computational aspects of the BHM modelling framework for NCD surveillance, while
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20 89 in Section 4 we run a simulation study to evaluate advantages and drawbacks of the
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22 90 approaches presented in detecting areas deviating from the expected trend. Finally,
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24 91 in Section 5 we conclude with a summary and some remaining discussion points.
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33 93 1.1 Data availability

34 94 One of the major challenges of surveillance studies is the availability of suitable data.
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36 95 This applies to both infectious disease and NCD surveillance. It is a particularly
37
38 96 important issue in low-income settings because surveillance studies often need to rely
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40 97 on information from surveys, and the lack of financial resources may make
41
42 98 comprehensive coverage of data sources (e.g. mortality/cancer registries) over an
43
44 99 entire country infeasible [9].
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47 100 In the last 15 years, a number of Health and Demographic Surveillance Systems
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49 101 (HDSS) have been established in low-income settings to provide a reliable source of
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51 102 health data and are now linked together through the International Network for the
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53 103 Continuous Demographic Evaluation of Populations and Their Health (INDEPTH, [10]).
54
55 104 Such continuous surveys are an invaluable source of data, but researchers face issues
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57 105 related to population representativeness. A recent study proposed a Bayesian
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3 106 probabilistic clustering method to evaluate the network representativeness in terms of
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5 107 socio-economic and environmental variables in sub-Saharan Africa, identifying areas
6
7 108 of poor coverage in the existing network and using predictive probability distributions
8
9 109 to suggest the best location for new HDSS sites [11].
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12 110 Even in high income settings where administrative resources are available for the
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14 111 entire population, there may be issues regarding the population at risk, used as
15
16 112 denominator in the risk estimates. In small-area studies on mortality or hospital
17
18 113 admissions, the denominator is usually the resident population in each administrative
19
20 114 area, typically estimated from national census statistics, but there may be estimation
21
22 115 problems for inter-censal years. In addition, it is not straightforward to define the
23
24 116 denominator where the interest is for less-defined geographies, such as the catchment
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26 117 areas of clinical centres (e.g. general practices in England).
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30 118 Furthermore, the availability of administrative or health data may become more limited:
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32 119 for instance, within the UK National Health Service, patients can now decide not to
33
34 120 share their medical records for research purposes. This clearly impacts on spatial
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36 121 coverage and could potentially lead to biased statistical inference if the data gaps are
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38 122 clustered in space and/or if they differentially affect specific population groups (e.g.
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40 123 elderly, more deprived) [12].
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45 46 47 125 1.2 Test-based methods

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49 126 Methods for NCD surveillance have largely been based around the idea of detecting
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51 127 whether the outcome of interest shows a particular behaviour in a defined subset (e.g.
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53 128 an area, period of time, or combination of space and time) when compared to the
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55 129 whole study region. Perhaps the most popular test-based methods used for NCD
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57 130 surveillance are the scan statistics. These were developed originally in the temporal
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3 131 setting only [13]; here, a fixed length “scanning window” is passed over the time-series
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5 132 data with the number of cases in the window being recorded. A log likelihood ratio
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7 133 (LLR) is calculated for each interval, and the test statistic is defined as the maximum
8
9 134 LLR over all intervals. This idea was extended to a spatial version of the scan statistic
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11 135 [23, 15], which was later further extended to the spatio-temporal setting [16]. In this
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13 136 case, the scanning window is represented by a cylinder, where the diameter specifies
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15 137 the spatial dimension and the height the temporal dimension. An additional version of
16
17 138 spatial scan statistic was proposed to account for correlation across spatial units which
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19 139 was not considered before [17]. Scan statistics have been extensively applied to
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21 140 numerous health care applications. Part of their popularity lies in the availability of free
22
23 141 user-friendly SaTScan™ software (<https://www.satscan.org/>). Recent applications of
24
25 142 SaTScan include the identification of signals for colorectal cancer [18], drug activity
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27 143 [19], criminality [20], and bat activity [21].
28
29 144 A further development has been the detection of spatial variations in temporal trends
30
31 145 (SVTT). These methods extend the scan statistics to estimate the time trend via a
32
33 146 regression-based model specifying either a linear or a quadratic function. The
34
35 147 quadratic SVTT method has, for example, been applied to cervical cancer data in
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37 148 women in the US from 1969 to 1995 highlighting areas where the risk was significantly
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39 149 different from the rest [22].
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41 150 Test-based methods such as scan statistics can only answer questions related to the
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43 151 deviation from the null hypothesis. An alternative approach is to explicitly model the
44
45 152 spatio-temporal structure in the data and to assess whether differences between
46
47 153 observed data and those predicted from the model provide evidence of anomalies.
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49 154 There are a number of advantages to adopting a model-based approach over a test-
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51 155 based method including the ability to: (i) have more statistical power to handle sparsity
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3 156 in the observed disease counts; (ii) explore more subtle departures from the
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5 157 expectation; (iii) account for the spatial and temporal correlation that is typically evident
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7 158 in health data; (iv) “borrow” information over space and time, therefore increasing the
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9 159 precision of the estimates generated; (v) include covariates that might explain some
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11 160 of the spatio-temporal variability.
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161 **2. Hierarchical models and likelihood-based inference**

162 Hierarchical models (HM) are able to deal with complex data structures, to exploit
163 dependencies between data sources and to propagate the inherent uncertainties that
164 are present in both the data and the modelling process. In the current context, an HM
165 combines two elements: a *process model* that describes how disease risk varies over
166 space and time, typically involving both extant covariate effects and a latent spatio-
167 temporal stochastic process; and a *data model* that describes the statistical properties
168 of the available health outcome data conditional on the realisation of the underlying
169 risk process. Both elements are specified up to the values of a set of unknown
170 *parameters*, which can be estimated by Bayesian or non-Bayesian versions of
171 likelihood-based inference, typically implemented using Markov chain Monte Carlo
172 integration and Monte Carlo likelihood maximisation methods, respectively. In addition
173 to estimating parameters, the scientific goals of health surveillance include *prediction*
174 of relevant properties of the unobserved risk surface as it evolves in real-time.
175 Parameters and latent stochastic processes are fundamentally different things, but
176 within the Bayesian paradigm they are both treated as unobserved sets of random
177 variables, and the operational calculus of estimation and prediction coalesces. In what
178 follows, we use BHM (Bayesian Hierarchical Modelling) as a shorthand for Bayesian
179 inference applied to a hierarchically specified model.

180

181 2.1 Space-time disease mapping

182 A class of BHMs that has been extensively used for the analysis of NCD data are the
183 so-called disease mapping models (DM). These are hierarchical models in which the
184 latent component of area-level disease risk is modelled as a spatially discrete Markov
185 random field [23] and, depending on the sampling design, the conditional distribution
186 of area-level case-counts is Poisson or Binomial [24, 25, 26, 27]. While the objectives
187 of these are descriptive, they have been used as the basis for the development of
188 detection models, which are framed in a surveillance perspective. Disease mapping
189 models have been extensively used to estimate and visualise the spatial or spatio-
190 temporal distribution of a disease (see for instance [9, 21, 28]).

191 Spatial dependence in the latent component of a DM is modelled by specifying
192 neighbourhood relationships amongst the area-level risks, the most widely used
193 definition being that two areas are *neighbours* if they share a common boundary. A
194 common choice for capturing temporal dependence is a random walk prior [29], but
195 extensions to incorporate spatio-temporal interactions among neighbouring areas and
196 time points have also been developed [30]. This framework can also account for
197 factors known to modify spatial and temporal trends that, in the context of NCDs, will
198 include demographic variables (e.g. age/sex/ethnicity) and social economic status.
199 Random effects can be assigned for each factor (with appropriate priors) and for
200 interactions if required. An example is provided by Goicoa *et al.* [31] who proposed a
201 space-time-age model to study prostate cancer incidence across 50 provinces in
202 Spain for 9 age groups over 25 years, accounting for all pair-wise interactions. The
203 authors used ranking of all provinces according to mortality rates to identify high-risk
204 groups.

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3 205 A key characteristic of BHMs is the ready availability of joint posterior/predictive
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5 206 distributions for parameters/latent processes and whatever of their properties are
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7 207 relevant to the public health questions of interest. In the context of disease mapping
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9 208 this leads to a spatial, temporal and spatio-temporal risk distribution that researchers
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11 209 can map, both in terms of point estimates but also of associated measures of
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13 210 uncertainty. For the latter, a common choice is the predictive probability that the
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15 211 relative risk exceeds a pre-specified threshold [32, 33]. Exceedance probabilities can
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17 212 be used to flag areas and/or time points characterised by increased risk that may then
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19 213 be further investigated. In this way, disease mapping, though not formally a
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21 214 surveillance method, can be used as a descriptive tool for the identification of areas
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23 215 and/or time periods with marked deviation from expectation. It is important to note that
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25 216 the strong smoothing effect of disease mapping models leads to conservative risk
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27 217 estimates, hence to a small number of false positive findings, at the expense of a low
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29 218 power for detecting high risk areas with low signal. To minimise that, an extensive
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31 219 simulation was run to find the best threshold on the exceedance probability scale to
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33 220 classify an area as high risk [33]. The authors showed that a good trade-off between
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35 221 false-positive and false-negative rates is achieved with a probability above 0.8 for a
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37 222 relative risk to be higher than 1, however this largely depends on the number of
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39 223 expected counts, the number of areas and time points, and the spatial risk [34].
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41 224 As an example of the typical disease mapping output, Figure 1 shows the incidence
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43 225 of malignant melanoma in males, at the census ward level in England and Wales over
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45 226 the period 1985-2009 from the Environmental and Health Atlas produced by the UK
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47 227 Small Area Health Statistics Unit (SAHSU) [35]; the map on the left presents the spatial
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49 228 distribution of the posterior relative risk mean estimates, while the map on the right
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3 229 plots the posterior probability that the corresponding relative risk is above 1, using the
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5 230 categorisation suggested in [33].
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8 231  Figure 1 here
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10 232 Disease mapping models can be extended to two or more outcomes that might share
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12 233 spatial (and temporal) patterns, for instance due to common risk factors. A joint model
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14 234 allows information to be borrowed across the outcomes, thus helping stabilise
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16 235 estimates, particularly when the outcomes are rare. The shared component model [36],
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18 236 originally developed for two diseases, includes a common component (likely to reflect
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20 237 common risk factors), and a disease-specific one, which can point towards specific
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22 238 risk factors otherwise masked in a single disease model. It was applied to male and
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24 239 female lung cancer [37] and later extended to jointly model multiple diseases [30, 38],
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26 240 with an application on oral cavity, oesophagus, larynx and lung cancers in males in
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28 241 the 544 districts of Germany from 1986 to 1990. Recently, it was further extended to
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30 242 jointly model age- and gender-specific diseases [39].
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34 243 An alternative multivariate specification considers spatial and temporal terms explicitly,
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36 244 modelling the correlation among the outcomes in space/time. As an example, road
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38 245 traffic accidents characterised by different severity were analysed over the period
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40 246 2005-2011 at the ward level in England while detection of high-risk areas was
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42 247 performed using exceedance probabilities of the area ranks based on accident rates
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44 248 [40].
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50 250 2.2 Space-time anomaly detection

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52 251 The standard disease mapping approach has been used informally to detect
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54 252 anomalies (unusual observations) in space and time, i.e. areas and/or time points with
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56 253 trends different to the expected ones, adding a space-time interaction parameter into
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3 254 the latent process [37]. The detection of anomalies may indicate the presence of an
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5 255 emerged localised risk factor, the impact of an intervention, or differences in the quality
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8 256 of data, such as misdiagnosis of a disease, and under- or over-reporting of cases.
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10 257 Mixture models have been proposed as a formal approach to anomaly detection. In
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12 258 particular, Abellan *et al.* [41], developed a BHM model (termed STmix) where a mixture
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14 259 of two normal distributions characterised by different variances is specified for the
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16 260 space-time interaction. Then, the interaction is used to classify areas as *common* and
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18 261 *unusual*. The authors performed a simulation study to compare the method against
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20 262 the standard disease mapping approach. The results of the simulation study showed
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22 263 that the standard approach was not able to capture the variability in the spatio-
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24 264 temporal interactions and therefore it was not able to distinguish between common
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26 265 and unusual areas. This is due to the excessive smoothing following the assumption
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28 266 of a common variance across all the areas and time points. STmix was applied to
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30 267 mammography screening data in Brisbane, Australia, at the statistical local area (SLA)
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32 268 level, from 1997 to 2008 in order to identify SLAs whose temporal trend exhibited
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34 269 volatility [42]. A well-known drawback of this approach is its limitation to incorporate
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36 270 specific time patterns, for example step changes that could signal the emergence of a
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38 271 new risk factor.
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44 272 Another mixture model, proposed by Li *et al.* [43], accommodates this issue. Here, the
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46 273 mixture specification of the method is defined directly on the relative risks in space
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48 274 and time, to allow for detection of areas with unusual time trends rather than space-
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50 275 time deviations. In particular, two alternative models are considered: the first one
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52 276 assumes a global time trend for all areas (*common trend*), while the second estimates
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54 277 a time trend for each area independently (*area-specific trend*). Through a simulation
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56 278 study the authors showed better performance in terms of both sensitivity and
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3 279 proportion of false positives compared to SaTScan on a wide range of scenarios. This
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5 280 approach, named BaySTDetect, was applied to detect unusual trends for asthma and
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7 281 chronic obstructive pulmonary disease at Clinical Commissioning Group (CCG) level
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9 282 in England (211 in total) on monthly data between August 2010 and March 2011
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11 283 across mortality, hospital admissions, and general practice drug prescriptions [44].
12
13 284 To illustrate the typical output obtained from this model, Figure 2 shows the area-
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15 285 specific time trends of the CCGs that were detected as unusual plotted against the
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17 286 national trend. Other applications of this method include burglary data [45], gray whale
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19 287 abundance [46] and mammography data [42].
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24 288 Figure 2 here

25
26 289 This method was further extended to increase its flexibility by accounting for different
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28 290 space-time patterns in the unusual observations, as well as by allowing for longer time
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30 291 series to be analysed. This improved method, termed FlexDetect, had a better
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32 292 performance when compared to the original method through an extensive simulation
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34 293 study [47].
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37 294 *Multiple testing*

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39 295 As surveillance studies involve evaluating trends for different health outcomes, many
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41 296 areas and different time periods at the same time, false detections are likely to occur
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43 297 by chance. Bonferroni correction has been extensively used in epidemiology to correct
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45 298 for multiple testing, particularly in omics studies [48, 49], but it is well known that this
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47 299 approach leads to conservative results. Benjamini and Hochberg first introduced an
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49 300 alternative index, the false discovery rate (FDR) [50], as the expected value of the rate
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51 301 of false positive findings among all rejected hypotheses and used it in a frequentist
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53 302 approach. The same method was suggested in the context of descriptive spatial
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55 303 epidemiology, to obtain areas characterised by a Standardised Mortality Ratio different
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3 304 from 1 [51]. Even in the Bayesian setting, FDR rules were suggested by many authors
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5 305 [52, 53, 54, 55, 56]. The mixture model proposed by Li et al. uses the specification
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7 306 suggested by Newton et al. [53] later used by Ventrucchi et al. [57] in order to account
8
9 307 for multiple testing. The authors base the FDR statistic on the posterior model
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11 308 probability, which represents the likelihood of the space-time unit investigated to follow
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13 309 the common trend model, i.e. to exhibit a risk pattern not deviating from the expected
14
15 310 one [43].

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19 311 While the importance of controlling for multiple testing is clear in classical significance
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21 312 testing, the analogous problem in predictive setting is less of a concern [58]. One
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23 313 reason for this is that local predictions from hierarchical models are naturally smoothed
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25 314 towards the global mean, making these consequently less prone to false positive
26
27 315 findings than unsmoothed area-by-area interval estimates. Another is that the Monte
28
29 316 Carlo sampling method allows the computation of whatever joint probability statements
30
31 317 are required. For example, if the public health question is whether current risk exceeds
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33 318 an agreed acceptable level in all areas that do, and in no areas that do not, meet a
34
35 319 particular criterion such as adherence to a particular advisory policy, the correct
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37 320 predictive probability to attach to this statement can be calculated.

321 **3. Computational aspects**

322 One of the biggest challenges researchers face when analysing large and complex
323 space-time datasets is their computational burden. This applies particularly in the
324 small-area context, where the number of space-time units investigated can vary
325 substantially depending on the chosen spatial and temporal resolution, from few
326 hundreds to hundreds of thousand units, particularly when several outcomes are jointly
327 analysed (for instance Foreman et al. [59] considered jointly deaths/age/sex specific
328 space-time trends in the U.S).

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3 329 The degree of complexity of the model (e.g. the number of parameters) also impacts
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5 330 on the computational burden, for instance in terms of convergence time when running
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7 331 MCMC simulations. Under the Bayesian paradigm, the choice of the prior will also
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9 332 influence convergence; an informative prior, assuming that there is no conflict with the
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11 333 data, will normally speed up convergence, while a vague prior will most likely lead to
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13 334 longer time to reach convergence. Finally, the choice of software used for the analysis
14
15 335 will affect the model running time. The user-friendly software BUGS (Bayesian
16
17 336 inference Using Gibbs Sampling) [60] has been traditionally used for Bayesian
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19 337 inference using MCMC methods, however it can be slow when high dimensional data
20
21 338 and/or complex models are used. Other MCMC-based methods, such as Stan [61]
22
23 339 and NIMBLE [62], are currently attracting attention due to their active development
24
25 340 community. An alternative way of dealing with computational limitations is to use
26
27 341 approximative methods; for instance INLA (integrated nested Laplace approximations)
28
29 342 [63] has been successfully used for running space-time disease mapping models (e.g.
30
31 343 [31, 64]); however, this method is somewhat less flexible than the aforementioned
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33 344 ones and as it relies on Normality of the latent process is not able to deal with mixture
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35 345 distributions.
36
37 346 Computationally intensive BHMs benefit from high-performance computing clusters to
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39 347 speed up computation times, however these are not necessarily required. For instance
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41 348 in [40] road traffic accidents data of different severity in England were analysed
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43 349 simultaneously at the ward level (~8000), over 9 years (for a total of around 150,000
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45 350 units and 32,000 parameters); the analysis was run in OpenBUGS and took 20 to 27hr
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47 351 on an Intel Core processor at 3.40 GHz with 8 Gbytes of random-access memory. On
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49 352 a much bigger scale, a US small area study considered more than half a million units
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3 353 and nearly 6,000 parameters [59]; the analysis was implemented in Stan using higher
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5 354 performance computing (HPC) clusters for faster calculations.
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9 355 **4. Simulation-based example**

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11 356 Considering the multitude of space-time methods available, as described above, it is
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13 357 important to formally evaluate their respective detection performance. In this paper we
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15 358 carried out a simulation study to formally evaluate the detection performance and
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17 359 compare DM, STmix and FlexDetect (see description of the models in Section 2.2).
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19 360 Following the design initially proposed by [43], and later used by [47], we used real
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21 361 asthma hospital episode statistics (HES) data to generate 50 simulated datasets. The
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23 362 asthma dataset was obtained from SAHSU, Imperial College London and consisted of
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25 363 disease counts across 211 Clinical Commissioning Groups (CCG) in England for 15
26
27 364 months, from January 2010 to March 2011. The 50 simulated datasets were generated
28
29 365 to closely resemble the patterns seen in the real dataset. A standard spatio-temporal
30
31 366 model [27] was first fitted to the real data and the obtained parameters were selected
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33 367 for the generation of the simulated data. We selected 15 areas to deviate from the
34
35 368 overall time trend over the last 5 time points. For these 15 areas, we selected the
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37 369 signal to be increased by $\log(2)$ for time points 3 and 10, and decreased by $\log(2)$ for
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39 370 time points 6, 12 and 15 out of the total 15 time points. In this way, we ensured that a
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41 371 realistic scenario was used (for more details see [47] Supplementary materials,
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43 372 Scenario 1). The R code used for the data simulation, together with the three models
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45 373 written in BUGS can be found on <https://github.com/aretib/bayesSTmodels.git>.
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47 374 The results are presented in Table 1 in terms of four different performance measures.
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49 375 We defined TP as the number of true positives, FP as false positives, TN as true
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51 376 negatives and FN as false negatives respectively. Sensitivity measures the ability of
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53 377 the model to correctly classify an unusual observation as such, defined as $TP/(TP+FN)$,
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3 378 and similarly specificity measures the ability of the model to correctly classify a
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5 379 common observation as such ($TN/(TN+FP)$). In addition, it is crucial to control the
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7 380 proportions of observations that are falsely classified as unusual (false discovery rate
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9 381 (FDR)) and common (false omission rate (FOR)) respectively; these should typically
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11 382 not exceed a value of 0.05, specified based on the standard p-value threshold. We
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13 383 consider (i) two different thresholds for DM: 0.8 as commonly used and previously
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15 384 described (DM1); a more conservative threshold of 0.9 (DM2), under the assumption
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17 385 that false positives are more important to minimise than false negatives, and (ii) two
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19 386 different rules for STmix as presented in the original paper: an area is modified if at
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21 387 least for one time point the space-time interaction has a probability greater than 0.8 to
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23 388 be above 1 (STmix1); an area is modified if for at least three time points the space-
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25 389 time interactions have an average probability greater than 0.8 to be above 1 (STmix2).
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31 Table 1 here
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33 391 As can be seen, the disease mapping approach using the standard threshold of 0.8
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35 392 on the posterior probability scale (DM1) shows the worst performance; as expected
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37 393 the method is able to detect nearly all unusual areas, with a sensitivity of 0.979;
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39 394 however, roughly 79% of the detected findings are not actually unusual (FDR = 0.785)
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41 395 (Table 1). Fixing a 0.9 threshold (DM2), FDR decreases, despite still being above the
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43 396 standard threshold of 0.05, while at the same time sensitivity also decreases (0.660)
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45 397 (Table 1).
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49 398 The two mixture models returned more comparable performances. STmix1 gave no
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51 399 false positive results (FDR = 0) and a sensitivity of 0.773, while for STmix2 sensitivity
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53 400 increased to 0.969, but at the same time a much higher proportion of false positive
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55 401 was detected (FDR = 0.220). The results of FlexDetect provided a balance between
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3 402 the two extremes, giving a proportion of FDR equal to 0.019 and a sensitivity of 0.796.
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5 403 In terms of specificity and FOR, both STmix and FlexDetect behave similarly.
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7 404 Differences across the competing models were observed in terms of computation time,
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9 405 an important factor in assessing their performance. All models were run in an Intel
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11 406 Xeon Core processor 3.40GHz with 125GB RAM. Each of the 50 simulations took on
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13 407 average 33.4 mins for models DM1 and DM2, 39.2 mins for models STmix 1 and
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15 408 STmix2 and 66.8 mins for FlexDetect. The simulation results suggest that using
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17 409 disease mapping (DM) for surveillance purposes is not appropriate and that one of the
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19 410 mixture models designed for detection should be used instead. Between STmix and
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21 411 FlexDetect it is worth mentioning that STmix can only identify areas where anomalies
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23 412 are present, and not the time points when these occur. In addition to this, its detection
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25 413 mechanism does not consider specific patterns in the time trends. These can be
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27 414 accommodated by FlexDetect, which however is more computationally intensive. Also
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29 415 note that mixture models notoriously have problems converging, suffering from issues
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31 416 such as label switching, which lead to multimodal posterior distributions. Both the
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33 417 detection methods deal with this through the modelling specification, such as [41]
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35 418 constraining the variance of the modified areas to be larger than that of unmodified
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37 419 areas, or through informative priors on the variances of the two components [43].
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420 5. Discussion

421 In this paper we have presented an overview of the main statistical methods for
422 disease surveillance in the context of NCDs, both from a test-based and model-based
423 perspective and with a particular focus on the BHM approach, which provides a flexible
424 framework to allow for complex data dependencies present in surveillance studies.
425 Through a simulation study we showed that disease mapping is not satisfactory when
426 looking for data anomalies, while the two methods based on mixture models provide

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3 427 a better compromise between detecting areas characterised by a deviation from the
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5 428 expected trend and limiting false positives. Note that our perspective is on methods
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7 429 that detect single areas, rather than clusters of adjacent spatial units. If the interest
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9 430 lays on detection in the presence of spatial proximity, recent methods have been
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11 431 developed to combine clustering with spatial smoothing, see for example [65] and [66].
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13 432 An interesting aspect of the general hierarchical framework presented is that it can
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15 433 easily incorporate forecasting of the disease risk, which is relevant in the context of
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17 434 epidemiological surveillance to evaluate the need for resources/policies/costs in
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19 435 specific areas and future time points. Some work in this area includes Foreman et al.,
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21 436 [59] who, using annual vital statistics for 1974-2011 at the US state spatial resolution,
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23 437 forecasted mortality up to 2024, while Ugarte et al. [67] used P-splines to forecast
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25 438 cancer mortality counts in Spanish regions for 2009-2011 using data from 1975-2008.
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27 439 Most of the work presented is based on routinely collected data for retrospective
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29 440 studies. However, there is increased importance of early warning detection, so that
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31 441 unusual behaviour can be detected at the earliest possible time. Syndromic data, such
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33 442 as primary care data, drug prescriptions, nurse calls and home visits, which are
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35 443 indicative of a potential anomaly, may provide an additional level of information leading
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37 444 to a detection event before the data aberration occurs [68]. Diggle et al. [69, 70]
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39 445 analysed NHS non-emergency telephone calls reporting symptoms of gastrointestinal
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41 446 diseases. The authors specified a spatio-temporal point process on the location and
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43 447 time of the individual calls and modelled the spatial and temporal dependency on the
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45 448 intensity of the process. They used exceedance probabilities to define maps of
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47 449 potential outbreaks. Another example can be found in Morrison et al. [71]
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49 450 who forecasted multiple measures of healthcare utilization (including physician visits
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51 451 and prescriptions of asthma medication) within British Columbia, Canada, where
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3 452 seasonal wildfires produce high levels of air pollution, significantly impacting
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5 453 population health. Here, the focus was on efficient, near real-time, computation which
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7 454 was achieved using INLA to perform approximate Bayesian inference.
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10 455 Potentially syndromic information can also be linked with routine data such as Hospital
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12 456 Episode Statistics and provide predictors in order to obtain a better description of the
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14 457 data and more accurate one-step-ahead forecasts. Lately work has been done to take
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16 458 advantage of the rich data from social media in a surveillance perspective. For
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18 459 instance, Dai et al. [72] linked tweets with the American Community Survey and the
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20 460 Behavioral Risk Factor Surveillance System to study asthma prevalence at the State
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22 461 level in the US. The authors claimed that the inclusion of social media data could be
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24 462 a cost-effective real time health detection system. However, there may be challenges
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26 463 in future due to selective data availability following perceived concerns about data
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28 464 security and confidentiality, as demonstrated by the newly implemented NHS National
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30 465 Data Optout Programme. This will potentially lead to bias in population
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32 466 representativeness due to non-random missingness [12] which will need to be
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34 467 addressed using advanced statistical methods, for instance through the integration of
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36 468 data from appropriate surveys / cohorts, as proposed in the context of residual
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38 469 confounding [73].
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44 470 An important issue with surveillance studies is that of the spatial resolution and the
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46 471 type of geographical areas considered; modifying these might lead to different results,
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48 472 as the spatial distribution of the outcome will depend on these choices. For instance,
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50 473 if within-area variability is substantial, results from statistical inference might suffer
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52 474 from false negative observations, as potentially high-risk places are aggregated with
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54 475 low-risk ones. The more spatial variability is present in the data, the more profound
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56 476 the potential impact of the modifiable areal unit (MAUP) [74, 75]. As MAUP depends
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3 477 on the level of aggregation, this issue has been linked to ecological bias [76] and the
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5 478 general suggestion in the scientific literature is to consider the finest spatial scale
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7 479 available. This can be particularly challenging for rare diseases where the number of
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9 480 cases at small-area level are very low. Furthermore, the choice of spatial resolution is
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11 481 mostly dependent on data availability and sparsity. BHMs have been suggested as a
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13 482 way to, at least partially, deal with MAUP. As there is an explicit relationship among
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15 483 areas globally and/or locally, through structured random effects, places belonging to
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17 484 a particularly small area can influence results for other areas, hence alleviating the
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19 485 MAUP problem [77].
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23
24 486 A key aspect of surveillance studies concerns how to communicate information to
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26 487 public health researchers and policy makers. This is particularly challenging as the
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28 488 statistical modelling of surveillance data becomes more sophisticated. In this context
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30 489 it is essential to develop user friendly tools such as atlases, web applications and
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32 490 reporting services that allow for data visualisations and easy implementation of the
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34 491 advanced methodologies. The Environment and Health Atlas for England and Wales
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36 492 [35] (typical output from the Atlas was presented in Figure 1) is an example of work in
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38 493 this direction, providing stakeholders and the general public with a collection of maps
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40 494 to inform on the spatial distribution of environmental factors and diseases. Through
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42 495 the exceedance probabilities, these maps give a perception of the uncertainty around
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44 496 the area level relative risks estimates.
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49 497 Web applications allow the ready implementation of statistical methods and perform
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51 498 complex data analyses, often through interactive data visualisations. These can be
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53 499 particularly useful for practitioners less skilful in statistical modelling and programming.
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55 500 As an example, the Rapid Inquiry Facility (RIF) which is currently being redeveloped
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57 501 within SAHSU, is designed to facilitate disease mapping and risk analysis studies and
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3 502 has been employed by more than 45 institutions in a number of countries [78]. A more
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5 503 recent example is the SpatialEpiApp that integrates two methods for disease mapping
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8 504 and cluster detection [79].
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10 505 To conclude, in this paper we presented a range of BHMs, which have proved to be
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12 506 useful for non-communicable disease surveillance. The choice of model should
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14 507 depend on various factors and most importantly on the objective of the study,
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16 508 characteristics of the data, and computational resources. It is commonly
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18 509 recommended to perform simulation studies based on the data in hand, to inform the
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20 510 model and to select detection rules that are most appropriate in each case.
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24 511 We believe that epidemiological surveillance will be at the centre of future
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26 512 methodological research to match the continuous increase in data availability, e.g.
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28 513 through social media; this will also open up issues related to data integration, selection
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30 514 bias and spatio-temporal misalignment. At the same time there will be the need to
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32 515 reduce the computational burden of increasingly complex models applied to large
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34 516 datasets, in order to provide timely results for decision making.
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For Review Only

	FDR	FOR	Sensitivity	Specificity
DM1	0.785 (0.773, 0.800)	0.002 (0.000, 0.006)	0.979 (0.933, 1.000)	0.722 (0.695, 0.744)
DM2	0.191 (0.100, 0.267)	0.026 (0.020, 0.030)	0.660 (0.600, 0.733)	0.987 (0.981, 0.995)
STmix1	0.000 (0.000, 0.000)	0.017 (0.010, 0.024)	0.773 (0.683, 0.867)	1.000 (1.000, 1.000)
STmix2	0.220 (0.167, 0.300)	0.002 (0.000, 0.005)	0.969 (0.933, 1.000)	0.978 (0.969, 0.985)
FlexDetect	0.019 (0.015, 0.031)	0.005 (0.004, 0.006)	0.796 (0.763, 0.827)	1.000 (0.999, 1.000)

Table1: Results of the simulation study to compare the detection performance of disease mapping (DM), the mixture model on the spatio-temporal interaction (STmix1, STmix2) and the mixture model on the spatio-temporal rates (FlexDetect).

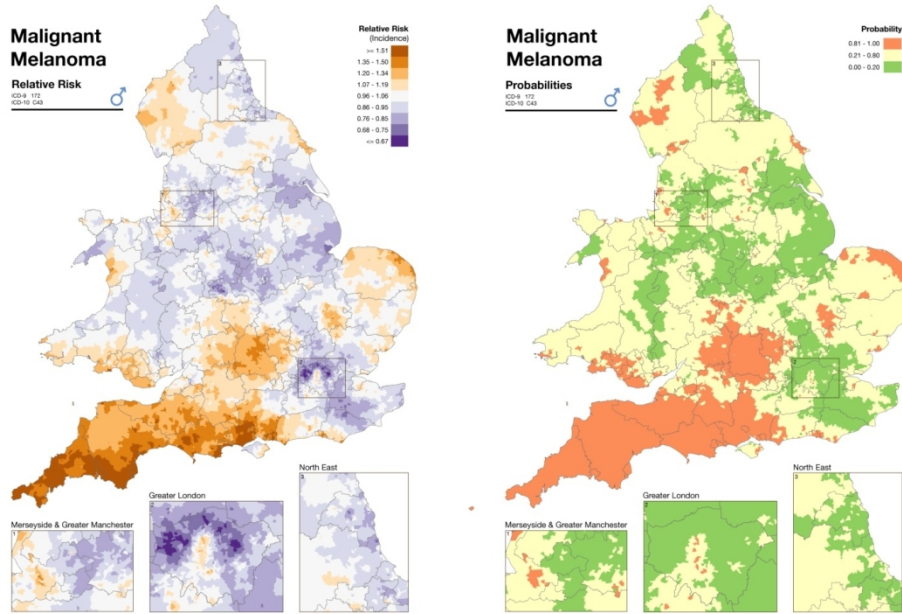


Figure 1: Area specific posterior mean relative risk of malignant melanoma (left) and posterior probability that an area is characterised by a relative risk above 1 (right). Source: Environment and Health Atlas [42].

548x386mm (72 x 72 DPI)

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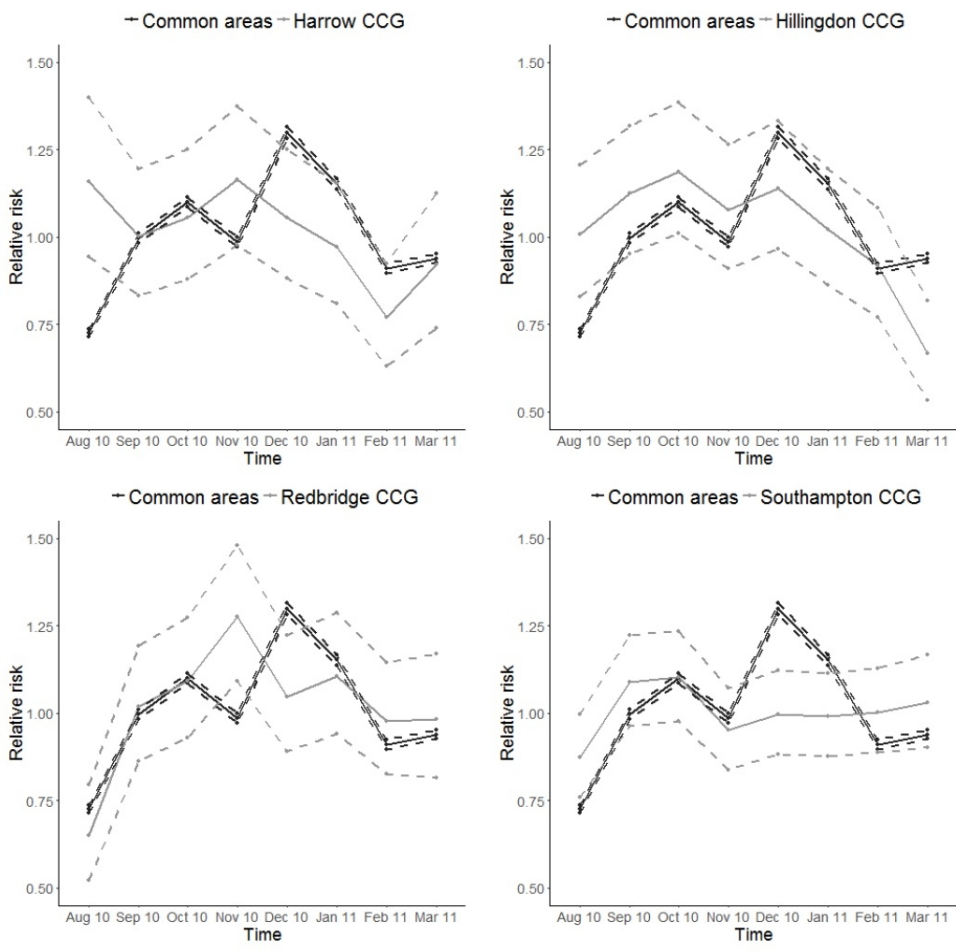


Figure 2: Relative risks and 95% credible intervals of hospital admissions for asthma and COPD for the national (common) temporal trend and for four areas classified as unusual.

377x371mm (72 x 72 DPI)