

1 **The Chord-Normalized Expected Species Shared (CNESS)-distance represents**  
2 **a superior measure of species turnover patterns**

3 **Running title:** *Measuring species turnover by CNESS*

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11

## 12 **Abstract**

- 13 1. Measures of  $\beta$ -diversity characterizing the difference in species composition  
14 between samples are commonly used in ecological studies. Nonetheless,  
15 commonly used dissimilarity measures require high sample completeness, or at  
16 least similar sample sizes between samples. In contrast, the Chord-Normalized  
17 Expected Species Shared (CNESS) dissimilarity measure calculates the  
18 probability of collecting the same set of species in random samples of a  
19 standardized size, and hence is not sensitive to completeness or size of  
20 compared samples. To date, this index has enjoyed limited use due to difficulties  
21 in its calculation and scarcity of studies systematically comparing it with other  
22 measures.
- 23 2. Here, we developed a novel R function **that** enables users to calculate ESS  
24 (Expected Species Shared)-**associated** measures. We evaluate the performance  
25 of the CNESS index based on simulated datasets of known species distribution  
26 structure, and compared CNESS with more widespread dissimilarity measures  
27 (Bray-Curtis index, Chao-Sørensen index, and proportionality based Euclidean  
28 distances) for varying sample completeness and sample sizes.
- 29 3. Simulation results indicated that for small sample size ( $m$ ) values, CNESS chiefly  
30 reflects similarities in dominant species, while selecting large  $m$  values  
31 emphasizes differences in the overall species assemblages. Permutation tests  
32 revealed that CNESS has a consistently low CV (coefficient of variation) even  
33 where sample completeness varies, while the Chao-Sørensen index has a high  
34 CV particularly for low sampling completeness. CNESS distances are also more  
35 robust than other indices with regards to undersampling, particularly when chiefly  
36 rare species are shared between two assemblages.

37 4. Our results emphasize the superiority of CNESS for comparisons of samples  
38 diverging in sample completeness and size, which is particular important in  
39 studies of highly mobile and species-rich taxa where sample completeness is  
40 often low. Via changes in the sample size parameter  $m$ , CNESS furthermore  
41 cannot only provide insights into the similarity of the overall distribution structure  
42 of shared species, but also into the differences in dominant and rare species,  
43 hence allowing additional, valuable insights beyond the capability of more  
44 widespread measures.

45

46 **Key words**

47  $\beta$ -diversity, CNESS, dissimilarity, species turnover, R function

## 48 Introduction

49 Reliable measurements of biodiversity are crucial in ecological studies, both with  
50 regards to species richness ( $\alpha$ -diversity) and assemblage composition ( $\beta$ -diversity).  
51 Whittaker (1960) defined  $\beta$ -diversity as the species turnover across spatial scale,  
52 with  $\alpha$ -diversity as the species richness at a sampling unit and  $\gamma$ -diversity as the total  
53 number of species over a large geographic area. Assessments of species turnover  
54 between samples as a key measure of  $\beta$ -diversity are commonly based on  
55 dissimilarity measures using mathematical descriptions of differences between pairs  
56 of samples (Legendre & Gallagher 2001; Tuomisto 2010; Mori, Isbell & Seidl 2018).  
57 These approaches are generally based on plot  $\times$  species matrices, often also  
58 including information on species' abundances, as basis for the calculation of the  
59 (dis)similarity or relative distance between pairs of samples.

60 The sampling effort for assemblages of diverse, mobile organisms, such as most  
61 insect assemblages, is difficult to standardize. The number of species in a sample  
62 generally correlates positively with the overall sample size and sampling effort, while  
63 sample completeness with regards to the local species pool is often unachievable in  
64 species-rich groups and biomes. Therefore, directly comparing the species records  
65 between two samples or sites with measures not accounting for the relative sampling  
66 effort and completeness creating a potential 'undersampling bias' that will result in  
67 highly unstable and unreliable outcomes (Coddington *et al.* 2009; Beck, Holloway &  
68 Schwanghart 2013; Iknayan *et al.* 2014). With regards to alpha-diversity,  
69 standardization can be achieved for example via the use of rarefaction (Hurlbert  
70 1971) and extrapolation techniques (Chao & Jost 2012; Chao *et al.* 2014), the use of  
71 species richness estimators (Hortal, Borges & Gaspar 2006) or by using parametric

72 diversity indices such as Fisher's  $\alpha$  (Beck & Schwanghart 2010). Nonetheless, most  
73 widespread measures of species turnover between assemblages are not  
74 appropriately accounting for the 'undersampling bias', with results potentially only  
75 **poorly** representing the "true" dissimilarity in the underlying populations (Beck,  
76 Holloway & Schwanghart 2013). For example, results are often heavily influenced by  
77 dominant species, or by widespread species of low abundance that by chance  
78 appear in only a subset of samples (Legendre & Gallagher 2001). Such problems  
79 are inherent in results gained by virtually all commonly used techniques to assess  
80 changes in species' assemblages, with incidence-based indices more sensitive to  
81 sample size than abundance-based ones (Beck, Holloway & Schwanghart 2013).  
82 **Some efforts have been made to address the influence of incomplete sampling on**  
83 **beta-diversity measures, both by developing indices regarded as less sensitive to**  
84 **sample size (Cardoso, Borges & Veech 2009; Schroeder & Jenkins 2018), or by**  
85 **trying to adjust existing indices (Chao *et al.* 2005; Yue & Clayton 2005) or using**  
86 **rarefaction techniques (Stier, Bolker & Osenberg 2016; Brocklehurst, Day & Fröbisch**  
87 **2018) that account for sample size-related variations in dissimilarity values. While**  
88 **these have yielded some interesting insights, they were often either plagued by very**  
89 **high levels of uncertainty or by low predictability power, making the interpretation of**  
90 **resulting values very difficult.**

91 One measure specifically designed to account for the issues relating to sample  
92 standardization is the 'Chord-Normalized Expected Species Shared' (CNESS)-  
93 distance. The CNESS index was introduced by Trueblood, Gallagher and Gould  
94 (1994), and it is based on the calculation of the 'Normalized Expected (number of)  
95 Species Shared' (NESS) between two samples as proposed by Grassle and Smith  
96 (1976). Both CNESS and NESS are in turn derived from the 'Expected Species

97 Shared' (ESS)-index that reflects the probability of obtaining the same set of species  
98 when randomly drawing a specific number of individuals from a community (Morisita  
99 1959; Grassle & Smith 1976). In other words, CNESS has been developed to cater  
100 for the effect that two samples of equal size randomly drawn from the same  
101 underlying community will by chance vary in their exact composition of species, and  
102 in the distribution of individuals across the different species. High CNESS  
103 dissimilarity values in this context reflect a low probability that two samples are  
104 drawn from the same community. Additionally, CNESS calculations allow for the  
105 sample size compared between two samples to be varied by adjusting the sample  
106 size parameter,  $m$ . This allows for a direct comparison of assemblages represented  
107 by two samples of varying sample size, by estimating their similarity for a  
108 standardized sample size common to both samples. In this context, small values of  
109  $m$  are believed to emphasize the similarity specifically in dominant species, whereas  
110 for large values of  $m$ , results are assumed to be increasingly affected by the  
111 composition of the entire species assemblage (Trueblood, Gallagher & Gould 1994).  
112 Calculating dissimilarities for different  $m$  values therefore generates unique insights  
113 into the similarity patterns between samples with regards to their different  
114 components (Trueblood, Gallagher & Gould 1994). CNESS has already been used  
115 particularly in studies of insect biodiversity, where samples are commonly showing  
116 large differences in the number of specimens caught at individual sampling events  
117 and in their sample completeness (Axmacher *et al.* 2004; Beck & Vun Khen 2007;  
118 Zou *et al.* 2014).

119 In spite of its theoretical advantages over other, commonly used dissimilarity metrics,  
120 the uptake of CNESS has been limited. For example, CNESS was excluded in a  
121 recent study by Schroeder and Jenkins (2018) who evaluated the sensitivity of

122 several dissimilarity indices to the ‘undersampling bias’, with the authors  
123 recommending measures such as the Bray-Curtis index due to their relative  
124 robustness to this effect. One of the reasons for the low profile of CNESS distances  
125 might relate to problems in calculating these dissimilarity values, with no suitable  
126 software tools available to date. The Compah96 software used in previous studies  
127 that is programmed in FORTRAN for MS DOS-based systems (Gallagher 1998) has  
128 become unavailable. In contrast, commonly used dissimilarity measures such as the  
129 Sørensen or Bray-Curtis indices can be calculated already by a number of standard  
130 packages in the open source R programming language (Oksanen *et al.* 2014). In  
131 addition, dissimilarity value of CNESS range between 0 and  $\sqrt{2}$  (see details in the  
132 method section), which makes direct comparisons with other dissimilarity measures  
133 whose values usually range between 0 (samples are the same) and 1 (samples are  
134 100% different) problematic.

135 Here, we provided scripts for a function to conveniently calculate the entire family of  
136 ESS (Expected Species Shared) measures using the R language (see Appendix 1)  
137 to make these dissimilarity measures more easily and widely available. We  
138 additionally introduced a slightly amended version of the CNESS measure adjusted  
139 so that values now range between 0 and 1. We used this function to explore how  
140 CNESS performs for assemblages of different species distribution structures for  
141 different sample size parameters,  $m$ . In addition, we evaluated the sensitivity of the  
142 CNESS measure in comparison to other, commonly used dissimilarity measures, with  
143 regards both to incomplete samples and variations in sample size. We used  
144 simulated rather than empirical data-sets to explore patterns and draw conclusions  
145 for the general behaviour of the different dissimilarity and distance measures.

146 **Method**

147 *The expression of CNESS*

148 CNESS is derived from the Expected Species Shared (ESS) measures introduced  
149 by Trueblood (Trueblood, Gallagher & Gould 1994). The ESS value for sites  $i$  and  $j$   
150 ( $ESS_{ij|m}$ ) (Grassle & Smith 1976), represents the number of species expected to be  
151 shared between two randomly selected samples of a standardized size of  $m$   
152 individuals, and can mathematically be expressed as:

$$ESS_{ij|m} = \sum_{k=1}^S \left[ 1 - \frac{\binom{N_{i^*} - N_{ik}}{m}}{\binom{N_{i^*}}{m}} \right] \times \left[ 1 - \frac{\binom{N_{j^*} - N_{jk}}{m}}{\binom{N_{j^*}}{m}} \right]$$

153 where  $S$  represents the total number of species,  $N_{i^*}$  and  $N_{j^*}$  represent the total  
154 number of individuals of site  $i$  and  $j$ , and  $N_{ik}$  and  $N_{jk}$  represent the abundance of the  
155  $k^{\text{th}}$  species at sites  $i$  and  $j$ .

156 While ESS calculations follow logical probability assumptions, the value of  $\binom{N_{i^*}}{m}$  for a  
157 large value of  $m$  can become almost infinite, leading to potential calculation failures  
158 during computation. The function nonetheless can be amended as follows (see  
159 mathematical proof in Appendix 2):

$$ESS_{ij|m} = \sum_{k=1}^S \left[ 1 - \prod_{n=0}^{m-1} \frac{(N_{i^*} - N_{ik} - n)}{N_{i^*} - n} \right] \times \left[ 1 - \prod_{n=0}^{m-1} \frac{(N_{j^*} - N_{jk} - n)}{N_{j^*} - n} \right]$$

160  
161 Although generally creating the same values for ESS, this formula is more robust  
162 with regards to the aforementioned calculation problems. The ESS values can in a  
163 next step be normalized, leading to the NESS (Normalized Expected Species



164 Shared) similarity measure between two samples, with values ranging between 0  
165 and 1 (Grassle & Smith 1976):

$$\text{NESS}_{ij|m} = \frac{2 \times \text{ESS}_{ij|m}}{\text{ESS}_{ii|m} + \text{ESS}_{jj|m}}$$

166

167 This measure is further modified to specifically account for the often large number of  
168 rare species that randomly occur in a small number of samples, even if samples are  
169 drawn from the same, underlying population. This modification is the CNESS (Chord-  
170 Normalized Expected Species Shared)-distance measure (Trueblood, Gallagher &  
171 Gould 1994). CNESS values can be calculated as:

$$\text{CNESS}_{ij|m} = \sqrt{2 \times \left[ 1 - \frac{\text{ESS}_{ij|m}}{\sqrt{\text{ESS}_{ii|m} \times \text{ESS}_{jj|m}}} \right]}$$

172 While NESS values vary between 0 and 1, Trueblood, Gallagher and Gould (1994)  
173 formulated CNESS in a way that theoretical values range between 0 and  $\sqrt{2}$ . This  
174 may result in difficulties when comparing its values with other dissimilarity indices  
175 that usually range between 0 and 1. We therefore slightly modified the CNESS index  
176 by removing the  $\sqrt{2}$  multiplier from the function, leading to the amended formula  
177 for CNESS<sub>a</sub>:

$$\text{CNESS}_{a(ij|m)} = \sqrt{1 - \frac{\text{ESS}_{ij|m}}{\sqrt{\text{ESS}_{ii|m} \times \text{ESS}_{jj|m}}}}$$

179 We have created an R function (Appendix 1) that conveniently allows us and our  
180 readers to calculate CNESS<sub>a</sub>, CNESS, NESS and ESS values in the R environment.  
181 The function contains three parameters, *x*, *m*, and *index* (by default, the *index* is set

182 as  $CNESS_a$ ); where  $x$  represents the species  $\times$  sample (as row  $\times$  column) matrix,  
183 and  $m$  the sample size parameter representing the number of individuals to be  
184 randomly drawn from the two samples that are compared. Theoretically, the choice  
185 of  $m$  can be any positive integer that is  $\geq 1$ . However, if the total sample size for a  
186 site is  $< m$ , this site will automatically be excluded from the analysis.

### 187 *Simulation and analysis*

188 To assess the performance of  $CNESS_a$  in comparison to other distance or  
189 dissimilarity measures, we first created a theoretical “control” dataset containing 100  
190 species. The abundance of these species was fitted to a logarithmic distribution  
191 pattern. The log-mean value of the resulting dataset is 6.5, with a log-sd value of 1,  
192 with the resulting dataset representing the trial community therefore containing about  
193 100,000 specimens distributed across the 100 species. For “treatments”, we created  
194 assemblages of equal size and distribution patterns, but with different amounts of  
195 “dominant” (D) and “rare” (R) species shared with the control. Each treatment  
196 contained three different populations, sharing 25%, 50% and 75% of their dominant  
197 (D) or rare (R) species with our “control”. Thus, we created a total of six “treatment”  
198 assemblages. The “dominant” species group shares the most abundant species from  
199 the control group. For example, the 25% dominant species (D25) group shares the  
200 25 species most abundant in the control group with that group, while randomizing  
201 their respective species rank order in the new group. The remaining 75 species in  
202 this second group are “new species” when compared with the control group.  
203 Likewise, the rare species assemblage shares the least abundant species with the  
204 control, with species ranks again randomized. The overall abundance distributions  
205 for different datasets are displayed in Appendix 3.

206 The actual analysis of index performances was separated into two parts. In the first  
207 part, the relative influence of abundant and rare species on the CNESS<sub>a</sub> calculated  
208 for different *m* values was evaluated. This was achieved by calculating the pairwise  
209 CNESS<sub>a</sub> value between the “control” and “treatment” datasets, with *m* values  
210 increasing from 1 to 100,000.

211 The second part of the analysis focuses on comparisons of the stability of distance  
212 or dissimilarity values of CNESS<sub>a</sub> and a selection of other, commonly used  
213 abundance-based dissimilarity indices. We selected three indices: i) the Bray-Curtis  
214 index, which is the most commonly used abundance-based dissimilarity index that  
215 has been argued to also be relatively robust with regards to undersampling bias  
216 (Schroeder & Jenkins 2018), ii) The Chao-Sørensen index, which is an abundance  
217 based form of the Sørensen index developed by Chao *et al.* (2005) in order to  
218 reduce the species distribution bias inherent in incidence-based indices, and iii)  
219 proportion-based Euclidean distances. For the CNESS<sub>a</sub> index, we selected *m* values  
220 of 1, 10, 100 and 1000.

221 We simulated two sampling strategies in order to investigate the effects of  
222 incompleteness of samples, and of unequal sample sizes. The first strategy was to  
223 have an equal sampling coverage for both “treatment” and “control” datasets, with  
224 the coverage varying between 0.01% (~10 individuals), 1%, 10% and 100% (all  
225 specimens present in the sample). Our sampling coverage refers to the number of  
226 individuals sampled from the overall pool, while we also calculated the sampling  
227 completeness that refers to the proportion of species sampled in comparison to the  
228 total number of species contained in the pool. Species completeness reach 9%, 54%  
229 and 97% for the individual coverage at 0.01%, 0.1% and 1, and reach 100% when

230 individual coverage is higher than 10%. The second strategy then compared the  
231 dissimilarity or distance between two samples that varied in their coverage, again  
232 with the coverage in the individual treatment samples varying from 0.01% to 1%, 10%  
233 and 100%, but using a constant number of 1000 specimens for the control treatment.  
234 We calculated the pairwise distance or dissimilarity values between the “control” and  
235 “treatment” samples from these combinations for all the above indices, carrying out  
236 permutations with 1000 iterations.

237 It need to be noticed that the main aim of our study was to test the ‘stability’ or  
238 ‘robustness’ of distance measures based on CNESS<sub>a</sub> and the other indices for  
239 differences in sampling coverage and unequal sample size scenarios, rather than  
240 evaluating how each index specifically reflects the underlying differences between  
241 samples and assemblages. The applicability of individual indices may partly depend  
242 on the actual sample patterns, with some measure comparisons in this regard  
243 provided in earlier studies (Chao *et al.* 2005; Beck, Holloway & Schwanghart 2013;  
244 Barwell, Isaac & Kunin 2015). In order to evaluate the stability of the different indices  
245 under the different sampling strategies, we then compared the coefficient of variation  
246 (CV = SD / mean) of the permutations results. In order to check the change of  
247 dissimilarity under different levels of sampling coverage, we computed the change  
248 rate ( $D_{c,n}$ ) between the undersampled dataset ( $D_n$ ) and the final, full sample dataset  
249 ( $D_1$ , i.e. representing either the full dataset in sampling approach 1, or the 1%  
250 control dataset in approach 2) using the formula:

$$D_{c,n} = \frac{|D_n - D_1|}{D_1}$$

251

252 All calculations were conducted in R V3.1.2 (R Core Team 2014), and we used the  
253 “CommEcol” package (Melo 2014) to calculate the Chao-Sørensen index, while the  
254 package “plyr” (Wickham 2011) was used for the data sorting during the simulation.  
255 The simulation scripts can be found in Appendix 4.

256

## 257 **Results**

258 CNESS<sub>a</sub> distances between control samples and samples taken from assemblages  
259 sharing rare species with the control were generally larger than distances between  
260 control samples and samples sharing dominant species with the control.

261 Nonetheless, the difference between these scenarios decreased with an increase in  
262 the sample size parameter  $m$ , with the shared rare species sample distances  
263 decreasing and the distances for samples sharing dominant species initially  
264 decreasing, but then increasing (Figure 1). For very large  $m$  – values, distances for  
265 “rare” and “dominant” treatments converged towards a common value, representing  
266 the value when ESS accounts for the actual number of species for site  $i$  (ESS<sub>ii</sub>) and  
267 site  $j$  (ESS<sub>jj</sub>), and the shared total number of species between two sites (ESS<sub>ij</sub>).

268 Comparisons of the different dissimilarity metrics show that the CV values generally  
269 increase with a decrease in sample coverage across all indices and for both, equal  
270 and unequal sampling strategies, as well as across both, the rare and the dominant  
271 shared species scenarios. Only the Bray-Curtis measures shows an exceptional  
272 peak in CV at a sampling coverage of 1% for the unequal sampling strategy (i.e. both  
273 samples have the same coverage). In all scenarios, the CV of CNESS<sub>a</sub>, Bray-Curtis

274 and proportion-based Euclidean distances never exceeded 0.1 (<0.05 for CNESS<sub>a</sub> in  
275 most cases), while the CV of Chao-Sørensen exceeded 0.1 in several scenarios, for  
276 example for samples sharing dominant species with the control for a coverage  
277 <0.1%, reaching a maximum value of 0.69 (Figure 2). With regards to variations in  
278 the standardized sample size in CNESS<sub>a</sub> ( $m$ ), an increase in its value resulted in a  
279 lower CV in the scenario of shared dominant species, but in a higher CV in the “rare  
280 species” shared scenario (Figure 2).

281 Where rare species were shared between control and treatment samples, CNESS<sub>a</sub>  
282 showed a stable performance across different sampling coverages and sampling  
283 strategies, as the change rate in comparison to the full coverage value never  
284 exceeded 0.1 and remained <0.05 for the majority of cases. In comparison, the  
285 changes of all other three indices exceeded 0.1 in some cases, for example in  
286 scenarios where sampling coverage <0.1% (Figure 3). Where dominant species  
287 were shared between control and treatment samples, all indices showed high  
288 change values >0.1 under a sampling coverage < 0.1% (this value could not be  
289 calculated for CNESS<sub>a</sub>  $m=1000$ ), except for Bray-Curtis distances under the  
290 unequal sampling strategy, but the change for this index exceed 0.1 when sampling  
291 coverage reached 10% and 0.1% (Figure 3).

## 292 **Discussion**

293 The R function we developed for this study and present in the appendix enables  
294 users to calculate the entire family of ESS-related distance measures. It allowed us  
295 to simulate and compare the performance of these widely neglected dissimilarity  
296 measures with more widespread measures for communities across a wide range of  
297 shared species and sample completeness scenarios. **The values of the amended**

298 CNESS<sub>a</sub> range from 0 to 1, which enables users to compare results directly with  
299 common dissimilarity measures. The sample size,  $m$ , which by default is set to 1,  
300 can be changed according to the users' requirements. In the simulation we selected  
301 a low sampling coverage of 0.01%, (~ 10 individuals) as our lower margin. This  
302 coverage, equivalent to 9% in species richness-based sample completeness, is  
303 much lower than that used in previous simulation studies dealing with the  
304 undersampling issue, for example ~ 40% by Brocklehurst, Day and Fröbisch (2018),  
305 ~ or 30% by Beck, Holloway and Schwanghart (2013). Such a low number of  
306 individuals in a sample is actually not uncommon in real-life arthropod studies (e.g.  
307 Beck & Kitching 2009; Duan *et al.* 2016), although we are commonly unable to  
308 assert the correct number of species in a sampling plot given the associated effort  
309 that would be required to completely sample such communities. This is also one  
310 reason that simulated groups with known species and abundance distributions were  
311 used in this study.

312 Our first simulation confirms that pair-wised results based on CNESS distances are  
313 strongly influenced by the distribution of shared species. Previous, empirical studies  
314 often calculated species turnover for different values of  $m$ , following the assumption  
315 that a smaller value of  $m$  emphasizes the similarity of samples with regards to their  
316 dominant species (Brehm, Homeier & Fiedler 2003; Axmacher *et al.* 2004; Hilt &  
317 Fiedler 2005), while here, we for the first time analyse in detail the implications of  
318 changes in its value across a wide variety of values up to the entire generated  
319 species pool. For a small sample size parameter  $m$ , CNESS distances between  
320 treatment assemblages sharing rare species with the control assemblage are much  
321 higher than the ones sharing dominant species with the control. This is reflecting the  
322 basic probability calculations on which the measure is based, since when taking a

323 relatively small sample (i.e. a small value of  $m$ ), the probability of any of the  $m$   
324 individuals belonging to shared species is higher when assemblages share their  
325 dominant species rather than their rare ones. Nonetheless, for large values of  $m$ ,  
326 CNESS approaches a constant value, i.e. the (chord-normalized) proportion of the  
327 shared number of species between two samples in both scenarios (shared rare or  
328 common species), explaining the convergence of CNESS dissimilarity values for  
329 large values of  $m$ . This confirms that for small  $m$  values, results chiefly reflect  
330 similarities in the dominant species (e.g. Hilt & Fiedler 2005), while for large  $m$ -  
331 values, the dissimilarity reflects the overall turnover between samples in their  
332 underlying species pool, irrespective of the abundance of the individual species  
333 within that pool. Altering values of  $m$  therefore enables researchers to shift the focus  
334 from the share of abundant species to the overall species pool. This ability in our  
335 view makes CNESS already a superior measure of species turnover patterns, since  
336 other, widespread beta-diversity indices only generate one fixed value that is  
337 strongly influenced by the underlying species abundance distribution pattern (Beck,  
338 Holloway & Schwanghart 2013).

339 Comparisons of the CV values confirms the robustness of the CNESS measure of  
340 compositional dissimilarity across a wide range of scenarios, including in cases  
341 where two communities share rare species. In contrast, the high variance particularly  
342 of the Chao-Sørensen dissimilarity measure for a low sampling completeness  
343 suggests that this index is not suitable to measure compositional dissimilarity in such  
344 scenario. Where communities share chiefly their dominant species, most indices  
345 show a high change ratio under a low sampling coverage, which means they all do  
346 not provide strong representations of the actual dissimilarity between the two  
347 samples. Nonetheless, even under this condition, CNESS still performs much better



348 than Chao-Sørensen and Euclidean distance measures. It needs to be noticed that  
349 the performance of Bray-Curtis is likely influenced by the sampling strategy, i.e. its  
350 changing ratio showed the same increasing trend with the decrease of sampling  
351 coverage under equal sampling strategy (i.e. undersampling for both assemblages,  
352 or only for one of the assemblages). However, Bray-Curtis reached a peak in  
353 similarity for the 1% coverage under the unequal sampling strategy, i.e. it behaves in  
354 a more unstable and unpredictable way across these scenarios when compared to  
355 CNESS that shows similar performances under the two sampling strategies.

356 In this study, we calculated CNESS for different sample size parameters  $m$  and three  
357 widespread beta diversity indices based on simulated datasets of known dissimilarity  
358 and using different sampling scenarios, to compare the difference between the  
359 different dissimilarity measures. It needs to be stressed that we did not assess how  
360 close resulting values were to the “true dissimilarity”. Instead, we focused on the  
361 variance and change ratio observed in the indices, since in the ordination or in other  
362 visualization approaches used to present the data, plots are commonly grouped by  
363 their relative distance or dissimilarity values. A robust prediction of dissimilarities  
364 across the different scenarios and under repeat extraction of random samples from  
365 underlying assemblages in this context is seen as an absolutely crucial basic  
366 criterion (Brehm & Fiedler 2004). In this regard, our results clearly emphasize the  
367 suitability and superiority of CNESS in samples of diverging sample sizes. The value  
368 of CNESS is sensitive to the distribution structure of shared species, which can be  
369 reflected by the changing of the sample size parameter  $m$ . While being highly useful  
370 in studying the compositional difference for overall species assemblage, in many  
371 real-life cases, setting  $m$  to large values comes at the cost of having to remove a  
372 number of samples whose overall sample sizes are smaller than  $m$ . Nonetheless, the

373 **CNESS** uniquely allows to address this problem via variations in the sample size  
374 parameter according to the respective underlying data structure of the samples that  
375 are being compared. We generally recommend researchers to calculate **the CNESS**  
376 **(or similar measures such as NESS) dissimilarity** for a number of different  $m$  values  
377 to obtain insights both into the share in dominant species and across the overall  
378 species pool (see e.g. Brehm, Homeier & Fiedler 2003; or Axmacher *et al.* 2004).

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### 382 **Authors' contributions**

383 YZ and JCA conceived the idea. YZ wrote the script and did the analysis. YZ and J  
384 JCA wrote the manuscript.

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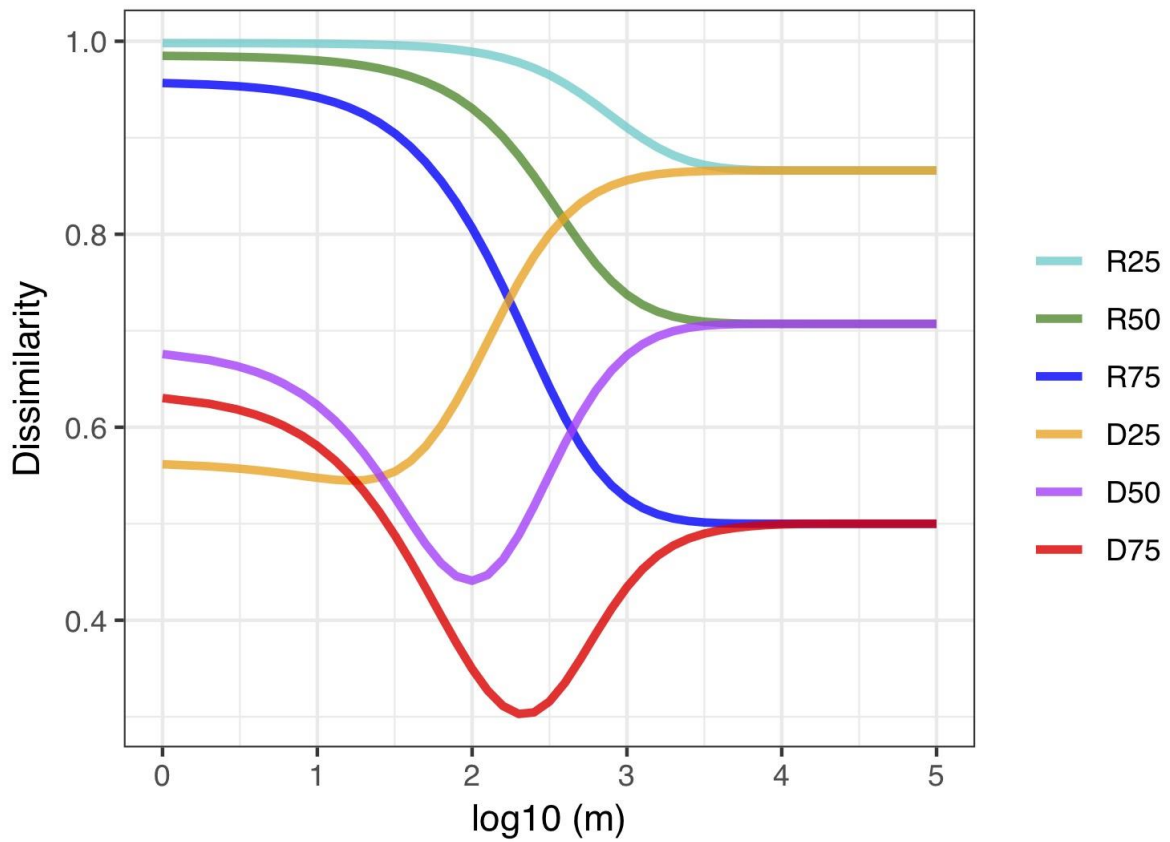
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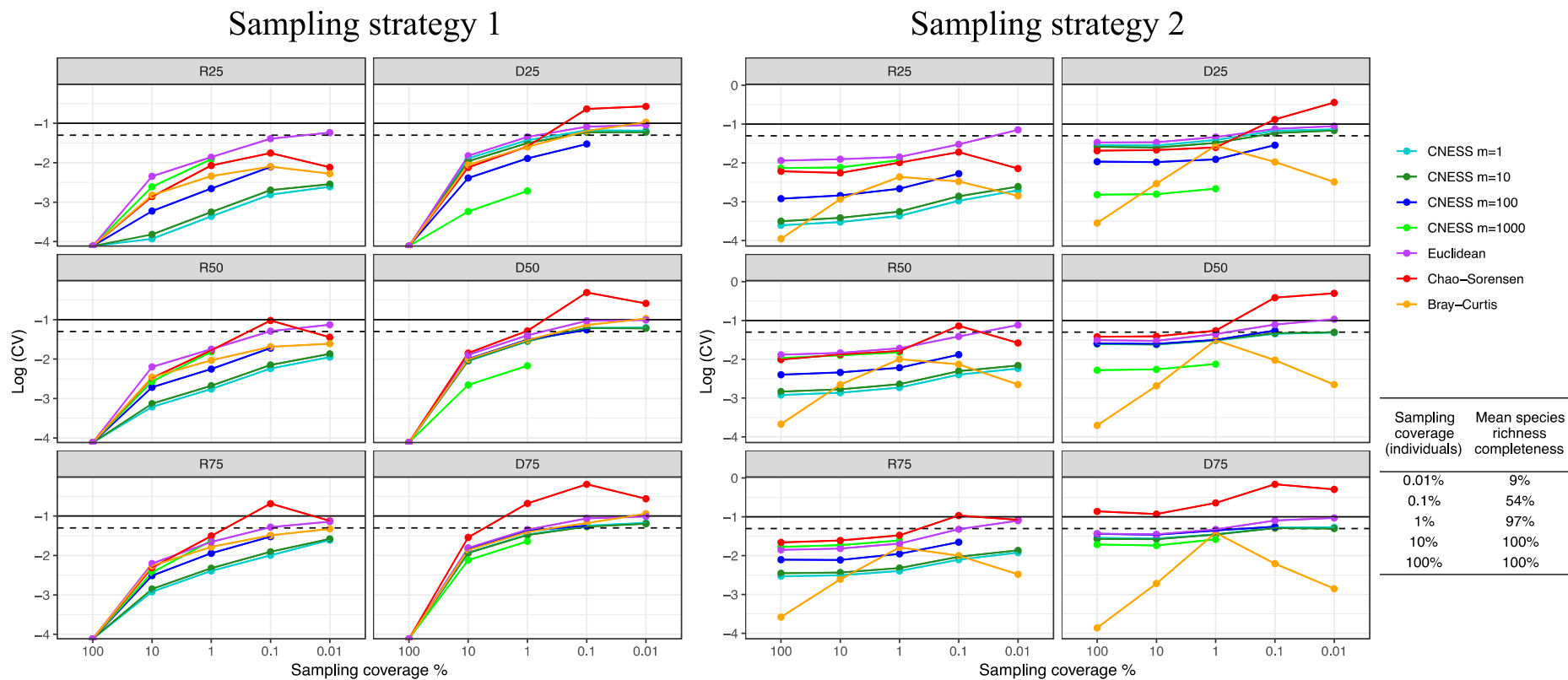
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474 **Figures**



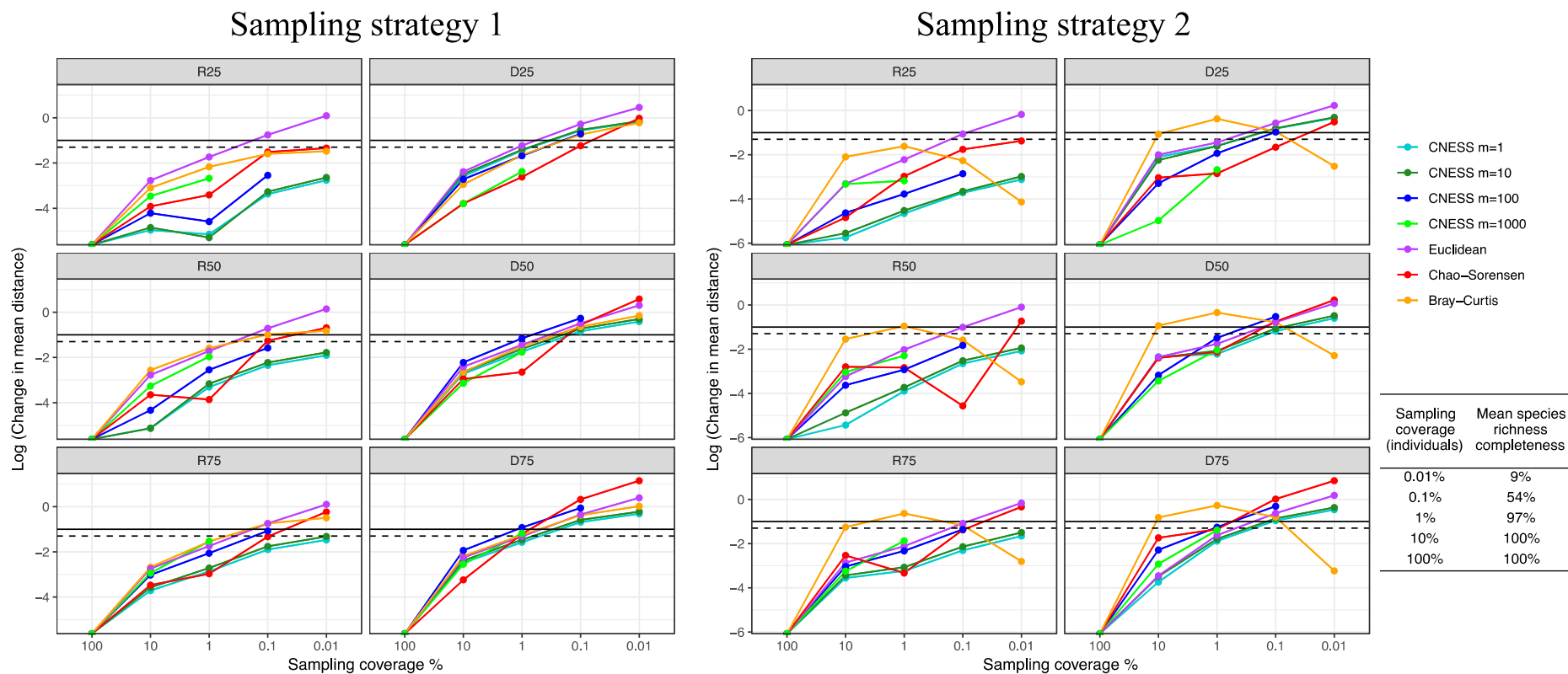
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476 Figure 1. The  $CNESS_a$  values calculated between the control and different  
477 treatment datasets with different  $m$  values. R25, R50 and R75 refer to treatments  
478 that share 25%, 50% and 75% of the rare species in the theoretical population, while  
479 D25, D50 and D75 refer to the respective share in dominant species with the control.



480

481 Figure 2. The coefficient of variation (CV, log<sub>10</sub> transformed) based on 1000 permutations for different dissimilarity or distance  
 482 measures calculated between the different treatments and the control sample for equal sampling (sampling strategy 1) and unequal  
 483 sampling (sampling strategy 2) for different sampling coverage. Solid and dashed vertical lines refer to 0.1 (log<sub>10</sub> value of -1) and  
 484 0.05 (log<sub>10</sub> value of -1.3) CV values. R25, R50 and R75 refer to treatments that share 25%, 50% and 75% of the rare species in  
 485 the theoretical population, while D25, D50 and D75 refer to the respective share in dominant species with the control. **The table**  
 486 **refers to the mean species richness completeness for different sampling coverages calculated based on the control group.**



487

488 Figure 3: Change in the mean value (log<sub>10</sub>-transformed) based on 1000 permutations for different indices between treatment and  
 489 control group for equal sampling (Sampling strategy 1) and unequal sampling (Sampling strategy 2) under different sampling  
 490 coverage. Solid and dashed vertical lines refer to 10% (log<sub>10</sub> value of -1) and 5% (log<sub>10</sub> value of -1.3) change. R25, R50 and R75  
 491 refer to treatments that share 25%, 50% and 75% of the rare species in the theoretical population, while D25, D50 and D75 refer to  
 492 the respective share in dominant species with the control. The table refers to the mean species richness completeness for different  
 493 sampling coverages calculated based on the control group.

494

495 **Electronic Supplementary Materials**

496 Appendix 1. R scripts to calculate Expected Species Shared (ESS) family

497 Appendix 2, Mathematical proof for the transformation of ESS formula

498 Appendix 3. The abundance distribution of simulated species for different “treatment”  
499 groups

500 Appendix 4. Simulation R scripts