

Possible Increase of Vegetation Exposure to Spring Frost under Climate Change in Switzerland

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Abstract: We assessed future changes in spring frost risk for the Aare river catchment that comprises the Swiss Plateau, the most important agricultural region of Switzerland. An ensemble of 15 bias-corrected regional climate model (RCM) simulations from the EXAR data set forced by the RCP 4.5 and RCP 8.5 concentration pathways were analysed for two future periods. Correlating actual meteorological observations and Swiss phenological spring index, we proposed and tested an RCM-compatible methodology (based on temperature data only) for estimating a start of spring and severity of frost events. In the historical climate, a significant advancement in start of spring was observed and frost events were more frequent in those years in which spring started sooner. In 2021–2050, spring is projected to start eight (twelve) days earlier, considering the RCP 4.5 (8.5) scenario. Substantial changes were simulated for the 2070–2099 period under RCP 8.5, when the total severity of frost events was projected to be increased by a factor of 2.1 compared to the historical climate. The study revealed the possible future increase of vegetation exposure to spring frost in Switzerland and that this phenomenon is noticeable even in the near future under the 'low concentration' RCP 4.5 scenario.

Keywords: spring frost; climate change; vegetation; Switzerland

1. Introduction

Late spring frost poses a major threat for vegetation's development and may result in considerable environmental and economic losses [1,2]. Recently, a warm start of spring 2017 was followed by an incursion of cold Arctic air into Western and Central Europe, which seriously harmed prematurely grown vegetation [3]. In general, the severity of such damage is linked to the lag between an onset of spring plant growth and a subsequent frost event [4].

Due to the ongoing climatic change and consequent rising temperatures, both start of growing season and frost-free period are being shifted towards the beginning of the calendar year. According to Jeong et al. [5], a growing season for temperate vegetation over the Northern Hemisphere has been starting earlier, the mean trend being about 2 days per decade in the 1982–2008 period. The prolonged growing season is linked to changes in the timing of plants' phenophases. For example, Menzel et al. [6] concluded that the average trend of spring/summer phenophase dates was −2.5 days per decade in Europe (phenophases tend to occur sooner). It should be noted, however, that these changes are species- and region-dependent [7,8]. According to Kolářová et al. [9], who analysed 18 common tree species in Central Europe, the largest advancements in spring phenophases were observed for shorter-lived, early-successional species (e.g., Prunus spinosa or Robinia pseudoacacia). The largest

shifts of phenophases towards the start of the year were observed in the northern parts of Central Europe, while advancements on the Mediterranean coastline were smallest (regarding apple trees) [10].

A key question in spring frost risk assessment is whether the advancement of phenophases is analogous to the shift of last spring frost. Wypych et al. [11] analysed trends of last spring frost dates over Central and Eastern Europe in the 1951–2010 period and showed that their magnitude is relatively variable over the domain. The western part of the area (roughly between 46–54 ◦N and 5–20 ◦E) had negative trends (last spring frost tends to occur sooner) of a magnitude ranging from –1 to −4 days per decade. This is in accordance with Bigler and Bugmann [12], who showed a negative trend of last spring frosts since the 1980s in Switzerland. By contrast, the trends in Eastern Europe (Ukraine, Belarus, western parts of Russia) were indistinct or even positive.

Many authors concluded that spring frost damage on vegetation has been increasing over middle latitudes in Northern hemisphere. Liu et al. [13] reported this phenomenon roughly across 43% of the hemisphere, especially in Europe. Kim et al. [14], using satellite data for the USA, showed that spring frost damage is linked to lower vegetation growth in spring and consequent lower vegetation greenness in summer. In addition, Augspurger [4] reported increased risk of spring frost damage, using 124 years of temperature records for the State of Illinois (USA). By contrast, Vitasse and Rebetez [3] concluded that the risk of damaging frost events to vegetation has remained unchanged in the 1864–2017 period, over the lowlands of Switzerland and Germany, due to comparable shifts in an onset of spring plant growth and late spring frosts, implying that changes in spring frost risk vary among regions and species analysed. This was shown by Vitasse et al. [15], who demonstrated that spring frost risk in Switzerland increased predominantly at stations located at elevations higher than 800 m a.s.l., while it remained mostly unchanged in lower altitudes.

Projections of changes in spring frost risk in a possible future climate are even more challenging. A probable decrease of spring frost risk was reported by Bennie et al. [16], who focused on deciduous trees (Betula pubescens) in Finland. Using climate change projections combined with phenological modelling, Molitor et al. [17] concluded that Luxembourg's winegrowing region will be less exposed to dangerous spring frost. By contrast, Leolini et al. [18] found an increased frequency of frost events at bud break in Central Europe for future scenarios. It should be noted, however, that these projections contain substantial uncertainties, most of which are related to the choice of climate model chains (greenhouse gas concentration scenario/global climate model/regional climate model/phenological model/estimated vegetation parameters [2]). This is in accordance with Mosadale et al. [19], who obtained opposite results when using different phenological models.

In this study, we endeavour to overcome uncertainties originating from a selection of climate model chains by developing a more robust approach. An onset of spring plant growth is estimated using temperature series only, which allows us to apply the definition within various data sets. The computed onsets of spring plant growth from temperature data only are evaluated against the Swiss spring index [20], calculated from actual phenological observations. The main aim of the study is to analyse changes in spring frost risk in a possible future climate of Switzerland. We focus on the Aare river catchment, which is probably the most vulnerable Swiss area in terms of late spring frost due to its agricultural importance. Moreover, it was one of many regions struck by a severe spring frost in April 2017 [3].

2. Data and Methods

2.1. Study Domain and EXAR Data Set

The analysis was performed for the Swiss Plateau and the northern slope of the Swiss Alps. The EXAR data set [21,22] used was originally produced in the context of a hydrological project 'Assessment of Extreme Flood Risks along the Aare and Rhine rivers' and refers to the Aare river catchment, covering an area of approximately 17,000 km^2 (Figure 1). The data set contains bias-corrected temperature and precipitation series at daily resolution (for both past climate and

future scenarios), representing averages for the domain. Thus, no spatial information is present in the data. The observational reference for a bias correction of reanalyses and climate models not the main term of our calibration of the corrective of the main corrections of the community contributions in $\frac{1}{n}$ is the CXAR data set was the daily temperature average of 11 stations (Chaumont, Neuchâtel, erven en en en een aanwere mar en aangeveerp van de avenge en een aanvers (commence), evenemies,
Château-d'Oex, Bern, Meiringen, Grimsel Hospiz, Andermatt, Altdorf, Luzern, Zurich, and Elm) from the MeteoSwiss network [23]. The bias correction was performed using the quantile mapping technique. In this approach, the climate model simulations are corrected to fit the cumulative distribution of The observational reference [24]. Each individual simulation is corrected separately. As model biases depend on the season [25], a time-dependent correction, calibrated for each calendar day using a 91-day moving window centred at the day of interest was applied. The day of interest was applied. Values of interest was appl moving window centred at the day of interest was applied. Values below the 1st percentile (or above moving window centred at the day of interest was applied. Values below the 1st percentile (or above 16 the 99th percentile) were corrected with the same correction as the 1st (99th) percentile. The cumulative the 99th percentile. distributions and correction functions for the 15 climate models' simulations (an example for the low distributions and correction functions for the 15 climate models' simulations (an example for the low concentration RCP 4.5 scenario) are given in Figure S1. The aforementioned observational reference for The aforementioned observational reference for bias correction was also used as a source of station bias correction was also used as a source of station data in our study. data in our study.

Figure 1. Location of study domain (red polygon) within Switzerland. All data used in this study **Figure 1.** Location of study domain (red polygon) within Switzerland. All data used in this study represent averages over this region. represent averages over this region.

Overall, 32 temperature series were obtained from the EXAR data set (1 representing station 1 reanalysis, and 30 climate model simulations). The station data series cover the 1900–2014 period. In data, 1 reanalysis, and 30 climate model simulations). The station data series cover the 1900–2014 order to extend the analysis back prior to the 20th century, the bias-corrected 20th Century Reanalysis period. In order to extend the analysis back prior to the 20th century, the bias-corrected 20th Century (version 2c, hereafter referred to as 20CRv2) [26] was used. Although 56 ensemble members of this reanalysis are available for the 1851–2011 period in EXAR, we used only its ensemble mean in the reanalysis are available for the 1851–2011 period in EXAR, we used only the cheemere mean in the study. For projections of future frost events, two ensembles of bias-corrected regional climate models ensemble means of future frost events, two ensembles of bias-corrected regional emitted frosted
(RCMs) enlisted in EXAR were used: (i) high-resolution EURO-CORDEX [27] RCM simulations (0.11° grid) driven by global climate models forced by the Representative Concentration Pathway (RCP) 4.5 scenario and ii) the same $RCM \times GCM$ ensemble forced by $RCP 8.5$. In both cases, the lateral boundary contains and H, are same RCM \sim GCM ensemble forced by RCP 6.6. In both case, are interest boundary conditions were taken from CMIP5 global climate model simulations [28], using observed greenhouse gas concentrations in the 1971–2005 period and the corresponding scenarios in the 2006–2099 period. Overall, 32 temperature series were obtained from the EXAR data set (1 representing station data,

gas concentrations in the 1971–2005 period and the corresponding securities in the 2006–2009 period.
The RCP 4.5 scenario represents stabilization of greenhouse gas concentrations without overshooting effective radiative forcing of 4.5 W/m^2 relative to pre-industrial values (~650 ppm) CO₂ equivalent). This is achieved by implementing mitigation policies [29]. By contrast, the RCP 8.5 $\frac{1}{2}$ contractive radiative forcing of the pre-industrial values ($\frac{1}{2}$). By contract, the KCP cost scenario represents a long-term large energy demand without implementation of mitigation policies, extrains represents a rong term ange energy demanta which imprementation or mingation policies, thus leading to high greenhouse gas emissions [30]. This scenario is presumed to reach an effective scenario representation representation representation of mitigation of mitigation policies, and mitigation policies, $\frac{1}{2}$ radiative forcing of 8.5 W/m² (~1370 ppm CO₂ equivalent). Individual ensemble members are listed in Table 1, and they are available for the 1971–2099 period.

Table 1. Regional climate models (RCMs) driven by global climate models (GCMs) from the EURO-CORDEX project extracted from the EXAR data set. All RCMs originally used the 0.11° rotated grid and are available driven by GCMs, forced by both RCP 4.5 and RCP 8.5 concentration pathways.

* Two individual runs of MPI-REMO-MPI simulations (r1i1p1 and r2i1p1) were used in the EXAR data set. Detailed information about individual RCMs' characteristics are provided in Kotlarski et al. [31].

2.2. Definition of Spring Frost

A key challenge is to determine whether a frost event occurred after an onset of spring plant growth, at which point plants become vulnerable to temperatures below 0 ◦C. Although it is possible to use observed vegetation phases from phenological data sets, this procedure cannot be applied into simulated data where the state of vegetation is simplified (in case of climate models) or contain many uncertainties (regarding variety of plant and soil physiological processes simulated by vegetation models). In order to overcome this issue, we proposed a methodology for estimating the beginning of an onset of spring plant growth (hereafter simply called 'start of spring') that is based solely on daily mean temperature (TAS) data.

The methodology is similar to a traditional growing degree days approach [32] but it was modified in several key points due to the nature of the study. Because our research domain (Figure 1) consists of various vegetation zones, it is impossible to use predefined species-dependent base temperatures and thermal times required for a calculation of growing degree days [33]. For each year (from January 1 to June 30), a cumulative sum function of TAS anomalies from a selected threshold was calculated. In contrast to a traditional growing degree days approach, negative TAS anomalies are also used (an example for three selected years is given in Figure 2A).

Figure 2. (A) Definition of start of spring; an example for years 1913, 1958, and 1990 using the 3.9 threshold. Solid curves represent cumulative sum functions, while black dots indicate starts of ◦C threshold. Solid curves represent cumulative sum functions, while black dots indicate starts of spring. (**B**) Correlation between starts of spring calculated from station data and the MeteoSwiss spring. (**B**) Correlation between starts of spring calculated from station data and the MeteoSwiss spring index using the 3.9 °C threshold. Dots indicate individual years, and the number 0.76 is the Pearson's correlation coefficient. (**C**) Differences in Pearson's correlation coefficients when using Pearson's correlation coefficient. (**C**) Differences in Pearson's correlation coefficients when using various threshold values. various threshold values.

In order to choose the most suitable TAS threshold for the whole domain, its selection was on a best fit of station data to the Swiss spring index (Figure 2B), which was calculated from in-situ based on a best fit of station data to the Swiss spring index (Figure 2B), which was calculated from observations of the Swiss Phenology Network [34] (see Brugnara et al. 2020). It is based on the blooming dates (hazel bush, coltsfoot, wood anemone, cherry tree, dandelion, lady's smock), new leaf blooming dates (hazel bush, coltsfoot, wood anemone, cherry tree, dandelion, lady's smock), new formation (horse chestnut tree, hazel bush, beech tree) and new needle formation of the larch [20].
The contract the larch state of the Fire mst principal component is then used to determine the dimensionless deviation of start of spring from the long-term mean. Positive values of spring index indicate that spring vegetation developed
https://www.https://www.https://www.https://www.https://www.https://www.https://www.https://www.https://www.https://www.htt developed based on the 1981–2010 average, while negative values show earlier vegetable development. TAS threshold values from 0.0 to $\frac{5.0}{\degree}$ C were tested for the overlapping 1951–2014 period, and the highest Pearson's correlation coefficient (0.76) was related to the 3.9 °C threshold threshold (Figure 2C). This threshold value was applied into all series analysed. The start of spring (Figure 2C). This threshold value was applied into all series analysed. The start of spring was defined as a date one day after the absolute minimum of the cumulative sum function (Figure 2A). Using the absolute minima of the cumulative sum functions, the method is robust against short term episodes of
The color color and the color $\frac{1}{2}$ above the selected threshold. In order to choose the most suitable TAS threshold for the whole domain, its selection was based The first principal component is then used to determine the dimensionless deviation of start of spring later than expected based on the 1981–2010 average, while negative values show earlier vegetation TAS above the selected threshold.

The severity of spring frost events was analysed using the frost index. The frost index was $\frac{1}{2}$ calculated for each year as a sum of negative TAS anomalies from 0 $°C$ (in absolute values) that occurred after the respective start of spring until June 30 (the rest of the year was excluded due to autumn at the start of spring until June 30 (the rest of the year was excluded due to autumn and winter frosts). Although frost resistance varies among individual species and depends on their
and winter frosts). Although frost resistance varies among individual species and depends on their phenological phase [2,35], the 0 °C TAS threshold serves as a natural limit for identifying severe spring frost (note that a TAS value of 0 °C may be linked to relatively harsh night-time frosts).

2.3. Other Characteristics and Statistical Testing 2.3. Other Characteristics and Statistical Testing

Statistical significance of estimated trends was assessed using the non-parametric Mann–Kendal Statistical significance of estimated trends was assessed using the non-parametric test. A two-sample Student's t-test was applied for analysing differences in the start of spring between years with spring frost and frost-free springs and for assessing significance of advancements in start of spring between the historical climate and future periods. Finally, a paired Student's t-test was used for analysing differences between spring dates in station data and 20CRv2.

3. Results 3. Results

3.1. Evaluation of EXAR Dataset 3.1. Evaluation of EXAR Dataset

In order to assess creditability of the EXAR dataset, seasonal TAS and its standard deviation In order to assess creditability of the EXAR dataset, seasonal TAS and its standard deviation from the bias-corrected 20CRv2 reanalysis and the CORDEX RCMs were evaluated against station from the bias-corrected 20CRv2 reanalysis and the CORDEX RCMs were evaluated against station data in the 1971-2000 period. The 15-member CORDEX ensemble performed well except for summer, when all simulations had a positive TAS bias ranging from 0.1 to 0.4 °C. For the 6 month period in the first half of the year (JFMAMJ, essential for the spring frost analysis), CORDEX tended to have a relatively small positive TAS bias, while a negative bias was found in 20CRv2. Overall, seasonal TAS was systematically lower in 20CRv2 compared to CORDEX (Figure 3A), which may be related to a combination of coarse resolution of original data and complex topography in the analysed domain (Section 2.1). CORDEX RCMs overestimated standard deviations of TAS (STDEV) in all seasons except summer, while standard deviation in 20CRv2 was smaller compared to station data through the year (especially in spring, Figure 3B). A higher (lower) value of STDEV in spring might be associated with more (less) severe frost events.

Figure 3. (**A**) Seasonal temperature bias for 15-member CORDEX ensemble (box plots) and 20th Century Reanalysis version 2c (circles) against station data. (**B**) Same as Figure 3A but for standard Century Reanalysis version 2c (circles) against station data. (**B**) Same as Figure 3A but for standard deviation of daily mean temperature. MAM stands for March 1–May 31 period, JJA for June 1–August deviation of daily mean temperature. MAM stands for March 1–May 31 persons for May 31 per in the Care of the M
MAM started for June 1–May 31 percent 1–May 31 percent 1–May 31 per in the Care of the Care of the Care of the 31, SON for September 1–November 30, DJF for December 1–February 28 (29) and JFMAMJ for January
1. J $J₁$ mam $J₂$ **Figure 3.** (**A**) Seasonal temperature bias for 15-member CORDEX ensemble (box plots) and 20th 1–June 30.

(Figure 3A) may result in earlier start of spring in CORDEX compared to station data. By contrast, the negative TAS bias found in 20CRv2 may result in later start of spring. The TAS bias was, however, relatively small (ranging from −0.3 to +0.3 °C), and modifying the threshold in accordance with bias values will not substantially improve the link between calculated starts of spring and spring indices in these data sets (considering correlation coefficients, Figure 2C). Moreover, it will have only a small effect on spring frost characteristics (Figure S2). \mathbf{N} is a small effect on spring from spring from spring from spring from spring from spring from spring \mathbf{N} . By applying the 3.9 °C threshold to all series used (Section 2.2), the overestimated TAS in JFMAMJ

3.2. Spring Frost in the Historical Climate 3.2. Spring Frost in the Historical Climate

In the station data (1900–2014), the start of spring tended to advance towards the beginning of In the station data (1900–2014), the start of spring tended to advance towards the beginning of the calendar year with a significant (at 1% level) trend of 1.8 days/decade. The earliest start of spring the calendar year with a significant (at 1% level) trend of 1.8 days/decade. The earliest start of spring (February 14) was recorded in 1990, while the latest end of winter was observed in 1919 when the (February 14) was recorded in 1990, while the latest end of winter was observed in 1919 when the spring began on May 2. Springs that started earlier were more prone to frost events, as spring frost spring began on May 2. Springs that started earlier were more prone to frost events, as spring frost was present during 37 out of 115 years (approximately every third year), and in those years, springs was present during 37 out of 115 years (approximately every third year), and in those years, springs started significantly (at 1% level) earlier compared to frost-free springs. The most severe spring frost started significantly (at 1% level) earlier compared to frost-free springs. The most severe spring frost (frost index = 12.2 °C) was observed in April 1913, approximately 1 month after the start of spring and an analogous event occurred in 1977 (Figure 4A). No clear trend of spring frost severity (analysed $\lim_{\alpha \to 0}$ index) was found in the station data.

Figure 4. Interannual variability in starts of spring (grey dots) with their least-squares trends (grey vertical state of \sim dashed lines) for station data (A) and the bias-corrected 20th Century Reanalysis, version 2c (B). Blue circles highlight springs with frost event(s) and their size indicates the value of frost index. Blue circles highlight springs with frost event(s) and their size indicates the value of frost index. Blue dashed lines represent periods between a start of spring and a last frost event.

In order to assess spring characteristics also in a period prior to the 20th century, a temperature series from 20CRv2 was analysed (Figure 4B). Two extremely early beginnings of spring were found in series from 20CRv2 was analysed (Figure 4B). Two extremely early beginnings of spring were found 1855 and 1872, and both of them were accompanied by a frost event. Overall, no significant differences in 1855 and 1872, and both of them were accompanied by a frost event. Overall, no significant in starts of spring between the station data and 20CRv2 were found, but 20CRv2 was unable to differences in starts of spring between the station data and 20CRv2 were found, but 20CRv2 was reproduce a significant advancement of spring starts towards the beginning of the year (considering unable to reproduce a significant advancement of spring starts towards the beginning of the year the overlapping 1900–2011 period). Analogously to the station data, springs that started earlier in the (considering the overlapping 1900–2011 period). Analogously to the station data, springs that started 20CRv2 were significantly (at 1% level) more prone to frost events but their severity was considerably early more eigenmentary (at 1% level) more prone to frost events but their severity mas constantinely smaller. This may be linked to the negative STDEV bias in 20CRv2 for the spring season (Figure 3B), bindicating smaller TAS fluctuations that are essential for frost events. In order to assess spring characteristics also in a period prior to the 20th century, a temperature

season (Figure 3B), indicating smaller TAS fluctuations that are essential for frost events. For an evaluation of CORDEX RCMs, spring characteristics in the relatively recent overlapping For an evaluation of CORDEX RCMs, spring characteristics in the relatively recent overlapping 1971 –2000 period were calculated. In this period, average dates for station data for start of spring in the station data for spring in the station of spring in the station data for station data for station data for stat versus 20CRv2 were March 20 and March 25, respectively. This relatively large difference (with versus 20CRv2 were March 20 and March 25, respectively. This relatively large difference (with respect1971–2000 period were calculated. In this period, average dates for start of spring in the station data

to no significant differences in starts of spring between the station data and 20CRv2 in the 1900–2011 period) is linked to much more pronounced advancement of the spring start in the station data (–9.3 days/decade) compared to 20CRv2 (–3.6 days/decade) in 1971–2000. The sum of the yearly frost indices for 1971–2000 (used for RCMs' evaluation) was 25.8 ◦C for station data and 9.0 ◦C for 20CRv2.

3.3. Future Changes of Spring Frost Risk

Because the springs that started earlier had significantly larger abundance of spring frost in the historical climate (Section 3.1), we tested continuity of this relationship in the warming world using simulated data. In the 1971–2000 period, the average date of spring start in the bias-corrected RCMs was in accordance with the station data and varied from March 17 (MPI-REMO-MPI-r2) to March 25 (SMHI-RCA-ICHEC, Table 2). Larger differences between RCMs were found in the sum of the yearly frost indices (ranging from 9.0 to 40.1 °C), but the ensemble mean was still close to the station data.

Table 2. Average start of spring (S) and sum of the yearly frost indices (If) in the bias-corrected regional climate models, station data and 20th Century Reanalysis (version 2c). Three time periods (1971–2000, 2021–2050 and 2070–2099) are considered. Values in parentheses indicate changes (ratios) in sum of frost index compared to the historical period in respective models. Bold values represent statistically significant change at 5% level and bold underlined ones at 1% level compared to the historical period (only starts of spring in individual models in future periods were tested).

In the near future (2021–2050), spring tends to begin earlier in 13 out of 15 RCMs considering the RCP 4.5 concentration pathway and in all RCMs under RCP 8.5. Those changes are particularly significant under RCP 8.5 (Table 2). On average, the start of spring advances by 8 days under the RCP 4.5 scenario and by 12 days considering RCP 8.5. Inter-model spread was larger compared to the historical period (Figure 5A), and one ensemble member (SMHI-RCA-MOHC forced by RCP 8.5) shifted the average start of spring even to the end of February. The overall advancement of spring's start was related to larger sums of the yearly frost indices (Figure 5B). In general, the sum of the yearly frost indices increased by a factor of 1.6 under RCP 8.5, and a slightly smaller increment (1.3) was simulated under RCP 4.5. This behaviour was model-dependent, however, because 5 out of 15 RCMs in both concentration scenarios simulated a decline of the frost index. In addition, the link between early (late) start of spring and high (low) frost index was not coherent in some RCMs. For example, the largest sum of the yearly frost indices (86.5 ◦C in IPSL-WRF-IPS under RCP 8.5) was associated with a rather typical average date for spring's start for the 2021–2050 period. Moreover, the average starts of spring in KNMI-RACMO-ICHEC were identical in both concentration pathways, but the yearly sums of the frost indices differed substantially (Table 2). It should be noted, however, that these discrepancies might be linked to data sampling, because a severe frost event may occur outside the 30-year-long analysed periods (e.g., an exceptional frost event in model year 2017 simulated by CLM-CCLM-MPI forced by RCP 4.5, Figure 6).

climate (grey) and two future time periods (2021–2050 and 2070–2099) using individual bias-corrected ensemble members. Green colour represents simulations forced by RCP 4.5, while red colour indicates simulations forced by RCP 8.5. Grey circle stands for station data, while grey square represents the bias-corrected 20th Century Reanalysis, version 2c. represents the bias-corrected 20th Century Reanalysis, version 2c. **Figure 5.** (**A**) Distributions of average start of spring and (**B**) sum of yearly frost indices in the historical

Additional shifts of the average start of spring towards the beginning of the year were simulated for the end of the 21st century (2070–2099). Under the RCP 4.5 concentration scenario, spring started nearly 3 weeks earlier on average compared to the historical period and a substantial statistically significant advancement in the start of spring was simulated under RCP 8.5 (Table 2, Figure 5A). At the end of the 21st century, many RCMs simulated springs that start right after the beginning of the year, meaning that winter (by our definition) ends before the close of the year or does not occur at all (Figures 6 and 7). Considering the RCP 8.5 scenario, all RCMs simulated higher values of the sums of the yearly frost indices with respect to their historical runs. The ensemble mean's sum of the yearly frost indices was larger by a factor of 2.9 compared to the station data and the increment was projected by 13 out of 15 RCMs (the two remaining RCM had values comparable to station data, Table 2). When forced by RCP 4.5, the RCMs simulated a smaller increase in the 2070–2099 period (2.0) compared to RCP 8.5.

Figure 6. Interannual variability of spring start dates (grey dots) in individual bias-corrected climate **Figure 6.** Interannual variability of spring start dates (grey dots) in individual bias-corrected climate model simulations forced by the RCP 4.5 concentration pathway. Blue circles highlight springs with model simulations forced by the RCP 4.5 concentration pathway. Blue circles highlight springs with frost event(s) and their size indicates the value of frost index. Blue dashed lines represent periods frost event(s) and their size indicates the value of frost index. Blue dashed lines represent periods between a start of spring and a last frost event. between a start of spring and a last frost event.

Figure 7. Same as Figure 6 but for simulations forced by the RCP 8.5 concentration pathway.

4. Discussion

Projecting changes in vegetation responses to a possible future climate is a challenging task that is accompanied by many uncertainties. The proposed method aims to provide a robust approach for assessing future spring frost risk for the complex Swiss terrain that contains several different vegetation zones, based on a widely available daily mean temperature only. The method was calibrated

in 1951–2014 using the Swiss spring index constructed from phenological in-situ observation [34]; however, it is possible that the relationship between temperature characteristics in spring and timing of vegetation's development would be changed in the future climate [36] because this link is probably not linear [37]. Many authors reported that climate change is responsible for decreased temperature sensibility of bud break. For example, Wang et al. [38] suggested that plants might be less likely to track climatic warming at locations with larger local spring temperature variance in order to avoid a frost risk and rely more on other cues, such as high chill requirements or photoperiods.

The 3.9 ℃ threshold value used in our method for identifying starts of spring corresponds to the growth-onset temperature identified by Breitenmoser et al. [39] for mid-latitudinal trees, however, it has to be considered that the threshold value represents an average over the whole analysed domain. Therefore, the threshold value has to be regarded with caution, because it may vary between vegetation zones [15]. Due to the size of the domain and its terrain heterogeneity, we employed daily mean temperature instead of commonly used daily minimum temperature when calculating the frost index. Low night-time temperature minima are often linked to geomorphologic features, such as altitude, shape of valleys, or exposition of slopes and may differ substantially between nearby locations [40,41]. The use of spatially more coherent daily mean temperature allows distinguishing of larger-scale frost events that are mostly related to an interruption of the prevailing westerly flow [42] and an incursion of a cold air mass from northern or eastern directions [43]. Cold northerly/easterly advection into Central and Western Europe is often linked to blocking anticyclones [44].

Improper simulation of frequency and persistence of blocking anticyclones is one of the largest drawbacks of current climate models [45,46], and it influences projections of circulation-induced temperature and precipitation extremes in a future climate [47]. Furthermore, Lhotka and Kyselý [48] showed that climate models tend to have too-cold northerly advection in winter, and this deficiency may propagate also into a spring season. Although the RCMs used simulated the spring frost index relatively well in the historical climate, an analysis of frost events' driving mechanisms is beyond the scope of this study. The 20CRv2 reanalysis did not capture the observed advancement of the spring start towards the beginning of the year, and the average date of spring start differed by 5 days compared to the observed data (considering the 1971–2000 period). Lorenz and Jacob [49] found weaker temperature trends in the NCEP/NCAR reanalysis over Europe compared to E-OBS and CRU observed data sets. Inasmuch as the 20CRv2 reanalysis shares many features with the NCEP/NCAR reanalysis [26,50], the erroneous, weaker temperature trend is probably present also in 20CRv2, and it is most likely related to this discrepancy.

Although substantial progress has been made in understanding relationships between increasing temperatures and shifts of phenophases, exact physiological mechanisms are still a subject of broad discussions. Cong et al. [51] reported that interannual variations in an onset of spring plant growth are extensively related to the number of chilling days over the Tibetan Plateau and suggested that continued future warming may lead to a deficiency in chilling and thus in changes in phenophase timing. An analogous phenomenon was reported in Switzerland by Asse et al. [52], who concluded that warmer winters significantly delayed bud burst and flowering along the elevation gradient. The role of a chill requirement, however, was questioned by Güsewell et al. [53], who found that reduced temperature sensitivity can result directly from spring warming alone. In addition, the importance of photoperiods for spring phenophases is not clear, and their roles are probably speciesand region-dependent [54,55]. Another important factor is a timing of snowmelt, especially in regions with higher elevation [56]. Finally, besides spring frost risk, other hazards such as heat waves and droughts or spread of pests [57] should be taken into account when preparing complex climate change adaptation and mitigation strategies.

5. Conclusions

Using 15 bias-corrected CORDEX RCMs from the EXAR data set, changes in characteristics of spring frost in Switzerland under climate change were analysed. The main conclusions can be summarized as follows:

- The methodology for estimating an onset of spring plant growth solely from temperature data was developed and tested. The estimated dates were well correlated (correlation coefficient = 0.76) with the Swiss spring index that was calculated from actual phenological observations.
- Significant (at 1% level) advancement of spring start was found in the observed data, with a trend of 1.8 days/decade in the 1900–2014 period. The 20CRv2 reanalysis failed to reproduce this trend.
- In the observed data, springs with early start were significantly (at 1% level) more prone to experience frost events compared to spring that began later. This relationship was, in general, simulated also in a future climate, but in some RCMs, a substantial spring advancement was not linked to a large sum of the yearly frost indices and vice versa.
- Considering the 2021–2050 period, spring is projected to start 8 or 12 days earlier (depending on concentration scenario). This advancement is linked to larger sums of the yearly frost indices compared to historical climate, approximately by a factor of 1.3 or 1.6, respectively.
- Major differences between concentration scenarios were found at the end of the 21st century (2070–2099). The earliest starts of spring and the largest values of the sum of yearly frost indices were simulated under RCP 8.5, in which the mean date of spring start was February 9 (about 6 weeks earlier) and the sum of the yearly frost indices was larger by a factor of 2.9 compared to historical climate.

Overall, we conclude that frost risk after an onset of spring plant growth will be larger in a future climate of Switzerland under both 'low' and 'high' concentration scenarios. The risk might be reduced by vegetation's adaptation mechanisms and changes in physiological processes, as discussed above, but it is questionable whether plants can adapt fast enough to follow the unprecedented (at least within the past several thousand years) warming trend projected under the RCP 8.5 concentration pathway. Further improvement may be associated with an analysis of regional patterns using the new CH2018 Swiss climate scenarios [58], which will allow for determining future frost risk on a 2×2 km grid.

Supplementary Materials: The following are available online at http://[www.mdpi.com](http://www.mdpi.com/2073-4433/11/4/391/s1)/2073-4433/11/4/391/s1, Figure S1: Cumulative distributions of daily mean temperature and correction functions for 15 climate model simulations and Figure S2: Interannual variability in starts of spring calculated for the bias-corrected 20th Century Reanalysis using two different thresholds

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