



Investigation of Laplace-Stieltjes transform for the ergodic distribution of the semi-markov random process with positive tendency, negative jumps and delaying boundary at zero

Tamilla Nasirova

Professor at Baku State University

nasirova.tamilla@mail.ru

Ulviyya Kerimova

Ph.D student at Institute of Cybernetics Azerbaijan National Academy of Sciences

ulviyye_kerimova@yahoo.com

Abstract: One of the important problems of stochastic processes theory is to define the Laplace-Stieltjes transform for the ergodic distribution of semi-markov random process. With this purpose, we will investigate the semi-markov random processes with positive tendency, negative jumps and delaying boundary at zero in this article. The Laplace transform on time, Laplace-Stieltjes transform on phase of the conditional and unconditional distributions and Laplace-Stieltjes transform of the ergodic distribution are defined. The characteristics of the ergodic distribution will be calculated on the basis of the final results.

Keywords: Laplace-Stieltjes transform; semi-markov random process; ergodic distribution; process with positive tendency and negative jumps.

2010 Mathematics Subject Classification: 60G05, 60J10, 60B05



Council for Innovative Research

Peer Review Research Publishing System

Journal: Journal of Advances in Mathematics

Vol 6, No. 3

editor@cirworld.com

www.cirworld.com, member.cirworld.com



Introduction

There are number of works devoted to definition of the distribution of the semi-markov processes and its main boundary functionals. Some authors are used the asymptotic, factorization and etc. methods ([2],[4],[5],[6],[9][12]) But other authors narrowing the class of distributions of walking are found the evident form for Laplace transforms for distributions and its main characteristics. In [7] The Laplace transformation for the distribution of the time of the system sojourn within a given band and its first and second moments are found .In [8] a model of inventory control is considered. It is described by a semi-markov random walk with a negative drift at an angle of $0 < \alpha < 90^\circ$, with positive random jumps, a delay, an absorbing screen at zero, and a reflecting screen for $a > 0$ at an angle α . The Laplace transformation is found for the distribution of the first moment storehouse exhaustion, and the first and the second moments are explicitly obtained. In [9] The Laplace-Stieltjes transform with respect to phase, the Laplace transform with respect to time, the conditional distribution, the unconditional distribution, and the Laplace-Stieltjes transform of the ergodic distribution of the process of semi-markov random walk with negative drift, nonnegative jumps, delays, and boundary screen at zero are obtained. In [10] The first passage of the zero level of the semi-markov process with positive tendency and negative jumps will be included as a random variable. The Laplace transform for the distribution of this random variable is defined. In [11] for the step process of semi-markov random walk with delaying boundary in $a > 0$ the evident form of Laplace transform by time was found.

The presented work explicitly defines the Laplace transform on time, Laplace-Stieltjes transform on phase of the conditional and unconditional distributions and Laplace-Stieltjes transform of the ergodic distribution for the semi-markov random processes with positive tendency, negative jumps and delaying boundary at zero.

1. Problem

Let's assume that in probability space $\{\Omega, F, P(\cdot)\}$ is given the sequence of independent, equally distributed and independent themselves positive random variables ξ_k and ζ_k , $k = \overline{1, \infty}$. Using these random variables we will derive the following semi-markov random process:

$$X_1(t) = z + t - \sum_{i=1}^{k-1} \zeta_i, \quad \text{if} \quad \sum_{i=1}^{k-1} \xi_i \leq t < \sum_{i=1}^k \xi_i, \quad k = \overline{1, \infty}$$

$X_1(t)$ is called semi-markov random processes with positive tendency and negative jumps.

General form of process semi-markov random walk with delaying boundary is given by A.A. Borovkov[1]

If process $X_1(t)$ is some process without boundary, then process $X(t)$ with delaying boundary at zero is defined following :

$$X(t) = X_1(t) - \inf_{0 \leq s \leq t} (0, X_1(s)) \quad \text{or} \quad X(t) = \max(0, \inf_{0 \leq s \leq t} (0, X_1(s)))$$

Idea of construction of the process semi-markov random walk is following :

Let $X_1(0) = z \geq 0$. Process $X(t)$ is equally to process $X_1(t)$ until, the process $X_1(t)$ is positive.

Let $X_1(t) \leq 0$; then $X(t)$ is equally to zero until, the process $X_1(t)$ will not have positive jump. In moment of jump of the process $X_1(t)$, process $X(t)$ will be have jump, such is equally to jump of the process $X_1(t)$.

The obtained process is called a process of a semi-markov random walk with positive tendency, negative jumps and delaying boundary at zero.

The aim of the present study is to find an evident form of the Laplace-Stieltjes transform of the ergodic distribution for $X(t)$.

2. Definition of Laplace transform on time for the distribution of the process $X(t)$

In accordance with formula of total probability for $x \geq 0$ we have

$$\begin{aligned} P\{X(t) < x \mid X(0) = z\} &= P\{X(t) < x; \xi_1 > t \mid X(0) = z\} + P\{X(t) < x; \xi_1 < t \mid X(0) = z\} = \\ &= P\{z + t < x; \xi_1 > t\} + \int_{s=0}^{\infty} \int_{y=0}^{\infty} P\{\xi_1 \in ds; X(s) \in dy \mid X(0) = z\} \cdot P\{X(t-s) < x \mid X(0) = y\} \end{aligned} \quad (1)$$

We denote



$$\left\{ \begin{array}{l} R(t, x | z) = P\{X(t) < x | X(0) = z\}, \quad x > 0 \\ \tilde{R}(\theta, x | z) = \int_{t=0}^{\infty} e^{-\theta t} R(t, x | z), \quad \theta > 0, \\ \tilde{\tilde{R}}(\theta, \alpha | z) = \int_0^{\infty} e^{-\alpha x} d_x \tilde{R}(t, x | z), \quad \alpha > 0. \end{array} \right.$$

(2)

In this case equation (1) will be as follows :

$$R(t, x | z) = P\{z + t - x < 0\}P\{\xi_1 > t\} + \int_{y=0}^t \int_{s=0}^t P\{\xi_1 \in ds\} d_y P\{\max(0, z + s - \zeta_1) < y\} R(t - s, x | y)$$

Both sides of this equation we applied Laplace transform by “t”

$$\int_0^{\infty} e^{-\theta t} R(t, x | z) dt = \int_0^{\infty} e^{-\theta t} \varepsilon(x - z - t) P\{\xi_1 > t\} dt + \int_{y=0}^{\infty} \tilde{R}(\theta, x | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{\max(0, z + t - \zeta_1) < y\} dP\{\xi_1 < t\}$$

where, $\theta > 0$.

Further

$$\tilde{R}(\theta, x | z) = \varepsilon(x - z) \int_0^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \int_{y=0}^{\infty} \tilde{R}(\theta, x | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{0 < y\} P\{z + t - \zeta_1 < y\} dP\{\xi_1 < t\}$$

After some simplifications we will get:

$$\begin{aligned} \tilde{R}(\theta, x | z) &= \varepsilon(x - z) \int_0^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \int_{y=0}^{\infty} \tilde{R}(\theta, x | y) \int_{t=0}^{\infty} e^{-\theta t} d_y \varepsilon(y) [1 - P\{\zeta_1 < z + t - y\}] dP\{\xi_1 < t\} = \\ &= \varepsilon(x - z) \int_0^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \tilde{R}(\theta, x | 0) \int_0^{\infty} e^{-\theta t} dP\{\xi_1 < t\} - \int_{y=0}^{\infty} \tilde{R}(\theta, x | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z + t - y\} dP\{\xi_1 < t\} = \\ &= \varepsilon(x - z) \int_0^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \tilde{R}(\theta, x | 0) \int_0^{\infty} e^{-\theta t} dP\{\xi_1 < t\} - \int_{y=0}^{\infty} \tilde{R}(\theta, x | y) \int_{t=\max(0, y-z)}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z + t - y\} dP\{\xi_1 < t\} \end{aligned}$$

If take into account

$$\max\{0, y - z\} = \begin{cases} 0, & \text{if } y < z \\ y - z, & \text{if } y > z \end{cases}$$

that is, why we get

$$\begin{aligned} \tilde{R}(\theta, x | z) &= \varepsilon(x - z) \int_0^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \\ &+ \tilde{R}(\theta, x | 0) \int_0^{\infty} e^{-\theta t} dP\{\xi_1 < t\} - \end{aligned}$$



$$\begin{aligned}
 & - \int_{y=0}^z \tilde{R}(\theta, x | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\} - \\
 & - \int_{y=z}^{\infty} \tilde{R}(\theta, x | y) \int_{t=y-z}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\}
 \end{aligned} \tag{3}$$

Both sides of this equation we applied Laplace transform by "x" [see (2)]

$$\begin{aligned}
 \tilde{\approx} R(\theta, \alpha | z) &= \int_{x=0}^{\infty} e^{-\alpha x} d_x \varepsilon(x-z) \int_{t=0}^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \\
 & + \tilde{\approx} R(\theta, \alpha | 0) \int_0^{\infty} e^{-\theta t} dP\{\xi_1 < t\} - \\
 & - \int_{y=0}^z \tilde{\approx} R(\theta, \alpha | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\} - \\
 & - \int_{y=z}^{\infty} \tilde{\approx} R(\theta, \alpha | y) \int_{t=y-z}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\}
 \end{aligned}$$

Take into consideration

$$\begin{aligned}
 \int_{x=0}^{\infty} e^{-\alpha x} d_x \varepsilon(x-z) \int_{t=0}^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt &= \int_{x=0}^{\infty} e^{-\alpha x} \varepsilon(x-z) d_x \int_{t=0}^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt + \int_{x=0}^{\infty} e^{-\alpha x} \int_{t=0}^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt d_x \varepsilon(x-z) \\
 &= \int_{x=z}^{\infty} e^{-\alpha x} d_x \int_{t=0}^{x-z} e^{-\theta t} P\{\xi_1 > t\} dt = \int_z^{\infty} e^{-\alpha x} e^{-\theta(x-z)} P\{\xi_1 > x-z\} dx
 \end{aligned}$$

At last we received the following integral equation for $\tilde{\approx} R(\theta, \alpha | z)$ when ξ_k and ζ_k , $k = \overline{1, \infty}$ equally distributed and independent themselves positive random variables

$$\begin{aligned}
 \tilde{\approx} R(\theta, \alpha | z) &= e^{\theta z} \int_z^{\infty} e^{-(\alpha+\theta)x} P\{\xi_1 > x-z\} dx + \\
 & + \tilde{\approx} R(\theta, \alpha | 0) \int_0^{\infty} e^{-\theta t} P\{\zeta_1 > z+t\} dP\{\xi_1 < t\} - \\
 & - \int_{y=0}^z \tilde{\approx} R(\theta, \alpha | y) \int_{t=0}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\} - \\
 & - \int_{y=z}^{\infty} \tilde{\approx} R(\theta, \alpha | y) \int_{t=y-z}^{\infty} e^{-\theta t} d_y P\{\zeta_1 < z+t-y\} dP\{\xi_1 < t\}
 \end{aligned} \tag{4}$$

We will solve this integral equation in special case.

Let's assume that ξ_1 random variable has the Erlangian distribution of n order, while ζ_1 random variable has the single order Erlangian distribution:



$$P\{\xi_1(\omega) < t\} = \left\{ 1 - \left[1 + \mu t + \frac{(\mu t)^2}{2!} + \dots + \frac{(\mu t)^{n-1}}{(n-1)!} \right] e^{-\mu t} \right\} \varepsilon(t), \quad \mu > 0$$

$$P\{\zeta_1(\omega) < t\} = [1 - e^{-\lambda t}] \varepsilon(t), \quad \lambda > 0$$

where

$$\varepsilon(t) = \begin{cases} 0, & t < 0 \\ 1, & t > 0 \end{cases}$$

In this case equation (4) will be as follows:

$$\begin{aligned} \tilde{R}(\theta, \alpha | z) &= \frac{(\alpha + \mu + \theta)^n - \mu^n}{(\alpha + \mu + \theta)^n (\alpha + \theta)} e^{-\alpha z} + \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n e^{-\lambda z} \tilde{R}(\theta, \alpha | 0) - \\ &- \lambda \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n e^{-\lambda z} \int_0^z e^{\lambda y} \tilde{R}(\theta, \alpha | y) dy - \\ &- \frac{\lambda \mu^n}{(n-1)!} e^{-\lambda z} \int_{y=z}^{\infty} e^{\lambda y} \tilde{R}(\theta, \alpha | y) \int_{t=y-z}^{\infty} e^{-(\lambda + \mu + \theta)t} t^{n-1} dt dy \end{aligned} \tag{5}$$

We will get differential equation from this integral equation. For this purpose, we will multiply both sides of equation (5) by $e^{\lambda z}$ and derive on z . Then we will multiply both sides of last equation by $e^{-(\lambda + \mu + \theta)z}$ and derive on z . If repeat this process $(n-1)$ time we have following differential equation:

$$\begin{aligned} \sum_{k=0}^n C_n^k \left[\lambda \tilde{R}^{(k)}(\theta, \alpha | z) + \tilde{R}^{(k+1)}(\theta, \alpha | z) \right] (-1)^{n-k} (\mu + \theta)^{n-k} + (-1)^{n-1} \lambda \mu^n \tilde{R}(\theta, \alpha | z) = \\ = (-1)^{n-1} \frac{[(\alpha + \mu + \theta)^n - \mu^n](\lambda - \alpha)}{(\alpha + \theta)} e^{-\alpha z} \end{aligned} \tag{6}$$

3. The general solution of the differential equation (6)

The general solution of this differential equation will be

$$\tilde{R}(\theta, \alpha | z) = C_1(\theta, \alpha) e^{k_1(\theta)z} + C_2(\theta, \alpha) e^{k_2(\theta)z} + \dots + C_n(\theta, \alpha) e^{k_n(\theta)z} + R_{sp}(\theta, \alpha | z) \tag{7}$$

where

$k_i(\theta)$, $i = 1, 2, \dots, n$, -are the roots of characteristic equation of (6)

$R_{sp}(\theta, \alpha | z)$ - is the special solution of the equation (5)

$$\tilde{R}_{sp}(\theta, \alpha | z) = A e^{-\alpha z}$$

where

$$A = \frac{(\lambda - \alpha)[(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) \prod_{i=1}^n [\alpha + k_i(\theta)]}$$

$$\left\{ \begin{aligned} \tilde{R}(\theta, \alpha | 0) &= \frac{(\alpha + \mu + \theta)^n - \mu^n}{(\alpha + \mu + \theta)^n (\alpha + \theta)} + \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n \tilde{R}(\theta, \alpha | 0) - \frac{\lambda \mu^n}{(n-1)!} \int_{y=0}^{\infty} e^{\lambda y} \tilde{R}(\theta, \alpha | y) \int_{t=y}^{\infty} e^{-(\lambda + \mu + \theta)t} t^{n-1} dt dy \\ \tilde{R}'(\theta, \alpha | 0) &= -\alpha \frac{(\alpha + \mu + \theta)^n - \mu^n}{(\alpha + \mu + \theta)^n (\alpha + \theta)} - \lambda \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n \tilde{R}(\theta, \alpha | 0) + \frac{\lambda^2 \mu^n}{(n-1)!} \int_{y=0}^{\infty} e^{\lambda y} \tilde{R}(\theta, \alpha | y) \int_{t=y}^{\infty} e^{-(\lambda + \mu + \theta)t} t^{n-1} dt dy - \\ &\quad - \frac{\lambda \mu^n}{(n-1)!} \int_0^{\infty} e^{-(\mu + \theta)y} \tilde{R}(\theta, \alpha | y) y^{n-1} dy \end{aligned} \right.$$

$$\sum_{k=0}^{n-1} C_n^k \left[\lambda \tilde{R}^{(k)}(\theta, \alpha | 0) + \tilde{R}^{(k+1)}(\theta, \alpha | 0) \right] = (-1)^{n-k} \frac{[(\alpha + \mu + \theta)^n - \mu^n](\lambda - \alpha)}{(\alpha + \theta)} - (-1)^{n-1} \lambda \mu^n \int_z^{\infty} e^{-(\mu + \theta)y} \tilde{R}(\theta, \alpha | y) dy$$

(8)

By finding $C_1(\theta, \alpha) \dots \dots C_n(\theta, \alpha)$ from equation (5) we will get the following system of algebraic equations:

By

exploitation of equation (7), equation (8) becomes

$$\left\{ \begin{aligned} \sum_{i=1}^n C_i(\theta, \alpha) + A &= \frac{(\alpha + \mu + \theta)^n - \mu^n}{(\alpha + \mu + \theta)^n (\alpha + \theta)} + \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n \left[\sum_{i=1}^n C_i(\theta, \alpha) + A \right] - \\ &\quad - \frac{\lambda \mu^n}{(n-1)!} \int_{y=0}^{\infty} e^{\lambda y} \left[\sum_{i=1}^n C_i(\theta, \alpha) e^{k_i(\theta)y} + A e^{-\alpha y} \right] \int_{t=y}^{\infty} e^{-(\lambda + \mu + \theta)t} t^{n-1} dt dy \\ \sum_{i=1}^n C_i(\theta, \alpha) k_i(\theta) - \alpha A &= -\alpha \frac{(\alpha + \mu + \theta)^n - \mu^n}{(\alpha + \mu + \theta)^n (\alpha + \theta)} - \lambda \left[\frac{\mu}{\lambda + \mu + \theta} \right]^n \left[\sum_{i=1}^n C_i(\theta, \alpha) + A \right] + \\ &\quad + \frac{\lambda^2 \mu^n}{(n-1)!} \int_{y=0}^{\infty} e^{\lambda y} \left[\sum_{i=1}^n C_i(\theta, \alpha) e^{k_i(\theta)y} + A e^{-\alpha y} \right] \int_{t=y}^{\infty} e^{-(\lambda + \mu + \theta)t} t^{n-1} dt dy - \\ &\quad - \frac{\lambda \mu^n}{(n-1)!} \int_0^{\infty} e^{-(\mu + \theta)y} \left[\sum_{i=1}^n C_i(\theta, \alpha) e^{k_i(\theta)y} + A e^{-\alpha y} \right] y^{n-1} dy \end{aligned} \right.$$

$$\sum_{k=0}^n C_n^k \left\{ \lambda \left[\sum_{i=1}^n k^n C_i(\theta, \alpha) + \alpha^n A \right] + \left[\sum_{i=1}^n k^{n+1} C_i(\theta, \alpha) + \alpha^{n+1} A \right] \right\} = (-1)^{n-1} \alpha^{n-1} \frac{[(\alpha + \mu + \theta)^n - \mu^n](\lambda - \alpha)}{(\alpha + \theta)} - (-1)^{n-1} \lambda \mu^n \int_z^{\infty} e^{-(\mu + \theta)y} \left[\sum_{i=1}^n C_i(\theta, \alpha) e^{k_i(\theta)y} + A e^{-\alpha y} \right] dy$$

(9)



Now we proof linear dependence of this algebraic system

If to consider the following substitutions:

$$\begin{aligned}
 \prod_{i=1}^n [\mu + \theta - k_i(\theta)] &= (-1)^{n+1} \lambda \mu^n \\
 \prod_{i=1}^n [\lambda + k_i(\theta)] &= (-1)^{n+1} \lambda \mu^n \\
 \prod_{i=1}^n [\mu + \theta - k_i(\theta)] &= (-1)^{n+1} [\lambda + k_j(\theta)] [\mu + \theta - k_j(\theta)]^n = (-1)^{n+1} \lambda \mu^n \\
 \prod_{i=1}^n [\alpha + k_i(\theta)] &= (\lambda - \alpha)(\mu + \theta + \alpha)^n - \lambda \mu^n \\
 \int_y^\infty t^{n-1} e^{-(\lambda+\mu+\theta)t} dt &= \frac{y^{n-1}}{\lambda + \mu + \theta} + \sum_{k=2}^n \frac{(n-1)!}{(n-k)! (\lambda + \mu + \theta)^k} y^{n-k} e^{-(\lambda+\mu+\theta)y}
 \end{aligned}
 \tag{10}$$

equation (9) becomes:

$$\begin{cases}
 \sum_{i=1}^n \left\{ [\mu + \theta - k_i(\theta)]^n - \mu^n \right\} C_i(\theta, \alpha) = \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} = \alpha \mu^n A \\
 \sum_{i=1}^n \left\{ [\mu + \theta - k_i(\theta)]^n - \mu^n \right\} C_i(\theta, \alpha) = \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} = \alpha \mu^n A \\
 \dots \\
 \sum_{i=1}^n \left\{ [\mu + \theta - k_i(\theta)]^n - \mu^n \right\} C_i(\theta, \alpha) = \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} = \alpha \mu^n A
 \end{cases}
 \tag{11}$$

Thus, (11) is a linear dependence equations system, as

$$C_2(\theta, \alpha) = C_3(\theta, \alpha) = \dots = C_n(\theta, \alpha) = 0$$

Then we have

$$C_1(\theta, \alpha) = \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) ([\mu + \theta - k_1(\theta)]^n - \mu^n) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} \tag{12}$$

Then the general solution of integral equation (5) will be as follows:

$$\begin{aligned}
 \tilde{R}(\theta, \alpha | z) &= \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) ([\mu + \theta - k_1(\theta)]^n - \mu^n) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} e^{k_1(\theta)z} + \\
 &\quad + \frac{(\lambda - \alpha) [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} e^{-\alpha z}
 \end{aligned}
 \tag{13}$$

This expression is the Laplace transform on time, Laplace-Stieltjes transform on phase for **conditional** distribution of the process $X(t)$

4. Ergodic distribution of the process.



We will need to find Laplace transform on time, Laplace-Stieltjes transform on phase for **unconditional** distribution of the process $X(t)$.

From construction process $X(t)$ is seen that

$$X(0) = X_1(0) = \xi_1(\omega)$$

Then we will get

$$\tilde{R}(\theta, \alpha) = \int_0^\infty \tilde{R}(\theta, \alpha | z) dP\{X(0) < z\}$$

Therefore

$$\tilde{R}(\theta, \alpha) = \int_{z=0}^\infty \left[\frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) ([\mu + \theta - k_1(\theta)]^n - \mu^n) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} e^{k_1(\theta)z} + \frac{(\lambda - \alpha)[(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} e^{-\alpha z} \right] d[1 - e^{-\mu z}]$$

or

$$\tilde{R}(\theta, \alpha) = \frac{\alpha \mu^n [(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) ([\mu + \theta - k_1(\theta)]^n - \mu^n) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} \frac{\mu}{[\mu - k_1(\theta)]} + \frac{(\lambda - \alpha)[(\alpha + \mu + \theta)^n - \mu^n]}{(\alpha + \theta) [\lambda \mu^n - (\lambda - \alpha)(\mu + \theta + \alpha)^n]} \frac{\mu}{[\alpha + \mu]}$$

This expression is Laplace transform on time, Laplace-Stieltjes transform on phase for **unconditional** distribution of the process $X(t)$.

Now, we will find Laplace-Stieltjes transform for ergodic distribution of the process $X(t)$.

In [3] (see p.363) proved a general theorem on the ergodicity of the process semi-markov random walk. The process described in this article a special case of this process.

Process $X(t)$ will be ergodic, if $E\xi_1 < E\zeta_1$, or

$$\frac{1}{\mu} < \frac{n}{\lambda} \Rightarrow \lambda < n\mu$$

If process $X(t)$ ergodic, then we can use Tauber's theorem [4]

$$Ee^{-\alpha X(\omega)} = \tilde{R}(\alpha) = \lim_{\theta \rightarrow 0} \theta \tilde{R}(\theta, \alpha)$$

We obtained

$$\tilde{R}(\alpha) = \frac{1}{(n-1)!} \frac{[(\alpha + \mu)^n - \mu^n] \mu}{[\lambda \mu^n + (\alpha - \lambda)(\alpha + \mu)^n]} \tag{14}$$

Expression (14) is Laplace-Stieltjes transform for ergodic distribution of the process $X(t)$. Respectively, we will get the following characteristics for $\lambda < n\mu$:



$$\tilde{R}'(0) = -EX(\omega) = \frac{n}{\lambda - n\mu}, \quad \lambda < n\mu$$

$$\tilde{R}''(0) - [\tilde{R}'(0)]^2 = DX(\omega) = \frac{2n}{(\lambda - n\mu)^2}, \quad \lambda < n\mu$$

5. Conclusions

In this article we have defined Laplace transforms on time, Laplace-Stieltjes transforms on phase for conditional and unconditional distributions and Laplace-Stieltjes transform for the ergodic distribution.

References

- [1] Borovkov A.A. *Veroyatnostnyye protsessy v teorii massovogo obsluzhivaniya (Probabilistic processes in queuing theory)*. Moscow: Nauka, 1972.
- [2] Borovkov A.A. On the asymptotic behavior of the distributions of first-passage. *Mat. Zametki*, 75(1):24–39, 2004.
- [3] Gikman I.I., Skorokhod A.V. *Theory of stochastic processes, Moscow, "Nauka", 1973*
- [4] Klimov G.P. *Stochastic service system, Moscow, "Nauka", 1966*
- [5] Lotov V.I. . On some boundary crossing problems for gaussian random walks. *The Annals of Probab*, 24:2154–2171, 1996.
- [6] Lebowitz J. L. and Percus J. K. *Asymptotic Behavior of the Radial Distribution Function* , J. Math. Phys. 4, 248 (1963)
- [7] Nasirova, T. I.; Sadikova, R. I. *Laplace transformation of the distribution of the time of system sojourns within a band* .Automatic Control and Computer Sciences (2009) 43: 190-194 , August 01, 2009
- [8] Nasirova, T. I.; Yapar, J.; Aliev, I. M. *A model of inventory control* Cybernetics and Systems Analysis (1999) 35: 553-562 , July 01, 1999
- [9] Nasirova T.H, Ibayev E.A., Aliyeva T.A. *The Laplace transform of the ergodic distribution of the process of semimarkovian random walk with negative drift, nonnegative jumps, delays and delaying screen at zero*. Theory of Stochastic Processes, 15(31),2009, №1, ISSN 0321-3900, pp 49-60.
- [10] Nasirova T.H, Kerimova U.Y. *Definition of Laplace ransform of the first passage of zero level of the semimarkov random process with positive tendency and negative jump*. Applied mathematics, 2011, 2, pp 908-912.
- [11] Omarova K.K *Laplace transformation of ergodic distribution of the step process of semi-markov random walk with delaying screen at positive point*. The Third International Conference "Problems of Cybernetics and Informatics" 2010.
- [12] Yu. V. Prokhorov *Asymptotic behavior of the binomial distribution* Uspekhi Mat. Nauk, 1953, Volume 8, Issue 3(55), Pages 135–142.
- [13] A. M. Walker *On the Asymptotic Behaviour of Posterior Distributions* Journal of the Royal Statistical Society. Series B (Methodological) (pp. 80-88) Vol. 31, No. 1, 1969.