International Journal of Modern Mathematical Sciences, 2014, 10(3): 239-246

International Journal of Modern Mathematical SciencesISSN:2166-286XJournal homepage:www.ModernScientificPress.com/Journals/ijmms.aspxFlorida, USA

Article

# On the Effect of Autocorrelation in Regression Model due to Specification Error

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*Article history*: Received 19 February 2014, Received in revised form 19 June 2014, Accepted 26 June 2014, Published 29 June 2014.

**Abstract:** The main cause of autocorrelation is omitted variables from the model. When an important independent variable is omitted from a model, its effect on the dependent variable becomes part of the error term. Hence, if the omitted variable has a positive or negative correlation with the dependent variable, it is likely to cause error terms that are positively or negative correlated. There are number of tests for specification error in detecting the errors of omitted variables from a regression analysis, one rarely knows the best test to use. This research uses bootstrapping experiment and some properties which estimators should possess if they are to be accepted as good and satisfactory estimates of the population parameters, the models investigated in the bootstrapping experiment consist of two autocorrelation models with autocorrelation level  $\rho = 0.5$  and 0.9. A bootstrap simulation approach was used to generate data for each of the models at different sample sizes (n) 20, 30, 50, and 80 respectively each with 100 replications(r). For the models considered, the experiment reveals that the estimated  $\beta$ 's were seriously affected by autocorrelation which may be due to omitted variables as the autocorrelation level varies in the different models (i.e. it produces a bias and inefficient estimator).

Keywords: specification error, autocorrelation, omitted variables, bootstrapping, estimators.

2000 Mathematical Subject Classification: 62J05

## 1. Introduction

Specification error is the error associated with the specification of the model, which can be take many forms such as omission of relevant variable, inclusion of unnecessary variables, errors of measurement etc.

When a relevant variable in the model is excluded, the specification will affect the properties of OLS estimator, in the presence of such error, OLS estimators will be bias. Some important questions that arise in the specification of the model include what variables should be included in the model, what are the probabilistic assumptions made about the  $y_t$  (dependent variable),  $x_t$  (independent variable) and  $u_t$ . (unobserved error).

Autocorrelation denotes the correlation between a time series  $y_t$  and its own lagged values  $y_{t-s}$ with  $s = -\infty \dots \infty$ .

$$y_t = \rho y_{t-1} + u_t$$

Or consider the error term in a linear

$$y_t = \beta X_t + \varepsilon_t$$
$$\varepsilon_t = \rho \varepsilon_{t-1} + u_t$$

An assumption in the linear regression model is that of zero value of the disturbance term covariance at all possible pairs of observation point. This is referred to as absence of autocorrelation of the error terms and when the disturbance term at any particular period is correlated with any other value of the disturbance term in the series, then we have autocorrelation. Some sources of autocorrelation include omitting explanatory variables, misspecification of the true random error. In many studies, the problem of autocorrelation has been subjected to both the theoretical and empirical investigation in the areas of investigating its consequences, detecting its presence and finding appropriate estimation methods. Test for detecting the presence of autocorrelation and alternative consistence methods of estimating linear methods with autocorrelated disturbance terms include the use of OLS and there are several approaches to resolving problems of autocorrelation, this include Lagging dependent variable, differencing the Dependent variable, GLS and ARIMA.

One common view of autocorrelation is that it is a technical violation of an OLS assumption that leads to incorrect estimates of the standard errors. From this perspective, analysts view autocorrelation as a potential sign of improper theoretical specification rather than a technical violation of an estimator assumption (Beck, 1985; Hendry and Mizon, 1978; Mizon, 1995). Regardless of which perspective on autocorrelation is adopted, lagged dependent variables have been proposed and utilized (for good or ill) as a solution. The effort of many researchers have been very rewarding like Cochrane and Orcutt (1949) brought to focus the attention of economist on the fact that the presence of autocorrelated errors terms require some modification of the usual least squares method of estimation. Bootstrap is a particular resampling scheme with replacement. In statistics and econometrics, bootstrapping has come to mean to resample repeatedly and randomly form an original initial sample using each bootstrapped sample to compute a statistic. The resulting empirical distribution of the statistic is then examined and interpreted as an approximation to the true sampling distribution.

#### 2. Methods

Considering the classical statistical linear model

$$Y = XB + U_t$$
 
$$U_t = \rho U_{t-1} + \varepsilon_t, \qquad var(\varepsilon_t) = \sigma^2$$
 
$$|\rho| < 1$$

$$\varepsilon_t \sim NID(0, \sigma^2), \qquad U_0 \sim N(0, \sigma^2_{\varepsilon_t})$$
  
where  $t = 1, 2, \dots, p$ 

Y = px1 vector observations on a sample space

X = pxk nonstochastic design matrix of rank k

 $\beta$  is a *k* –dimensional fixed vector of unknown parameters.

U = px1 vector of unobservable random variable with (0,1) and a finite covariance matrix.

Consider a two model of the form

Model
 Specification
 Problem

 I. True: 
$$y_t = 10.0 + 5.0x_{1t} - 2.0x_{2t} + u_t$$
 Null:  $y_t = \beta_0 + \beta_1 x_{1t} + \beta_2 x_{2t} + 0.5u_{t-1} + v_t$  Autocorrelation

 II.
 True:  $y_t = 10.0 + 5.0x_{1t} - 2.0x_{2t} + u_t$ 

 Null:  $y_t = \beta_0 + \beta_1 x_{1t} + \beta_2 x_{2t} + 0.9u_{t-1} + v_t$ 
 Autocorrelation

Where the true model is specifying a correct without autocorrelation problem and the null model is specifying the model that has been affected by autocorrelation.

Observations on the dependent variables are generated according to one of the specification labeled true.

#### **Generation of data**

For the autocorrelation data, first obtain the autoregressive error term  $\rho u_{t-1} + v_t (\rho = 0.5 \text{ and } 0.9)$  by generating 21 normal random deviates using Rand () command in Microsoft Excel, I standardized the series obtained. Calculate starting value  $v_0$  by drawing a random value U from N(0,1) and divide by  $\sqrt{1} - \rho^2$ . Use successive value of  $v_t$  drawn and initial value of  $v_0$  to calculate  $\rho u_{t-1} + v_t$  ignoring the first value in the series to avoid problem of initial value. The process above is repeated in as many times to obtain 100 replicationseach of series 20,30,50, and 80 respectively. After generating the autoregressive error term, the experiment is then repeated for each of the sample sizes 20,30,50, and 80 for the generation of the dependent variable.

For the bootstrap experiment, we consider the specification labeled True model:

 $y_t = 10.0 + 5.0x_{1t} - 2.0x_{2t} + u_t$ , assign a numerical values to all the parameter ( $\beta_0 = 10, \beta_1 = 5.0, \beta_2 = -2.0$  in the model, the variance  $\sigma^2$  is also assigned a numerical value on the basis of assumed $\sigma^2$ , and then the disturbance term *U* is generated. The *U* generated was standardized. A random sample of *X* was then selected from a pool of random numbers and numerical values of  $y_t = 10.0 + 5.0x_{1t} - 2.0x_{2t} + u_t$  was computed for each sample size using Microsoft Excel software. The *X*'s and *Y*'s generated were then copied from Microsoft Excel into STATA and then bootstrapped and replicated 100 times using a STATA command, each replication produces a bootstrap sample regression *Y* on fixed *X*'s. The procedure above is then repeated for different sample sizes and was also performed on each of our two models.

#### Criteria for evaluating the performance of the estimators

In this study, the following criteria were used for comparing evaluation of the performance of our estimators;

Average or mean of estimators in comparison with the true parameter, let  $\hat{\beta}$  be the average estimates of the parameter  $\beta$  obtained in the *i*<sup>th</sup> bootstrap replication, we compute

$$\hat{\beta} = \frac{\sum_{i}^{r} \beta}{r} , \qquad \text{where } r = \text{number of replications}$$
  
•  $\text{Bias}(\hat{\beta}) = \hat{\beta} - \beta$   
•  $\text{Variance}(\hat{\beta}) = \frac{1}{r} \sum_{i=1}^{r} (\hat{\beta} - \beta)^2$ 

$$MSE(\hat{\beta}) = \frac{1}{r}E(\hat{\beta} - \beta)^2$$

## 3. Results and Discussion

In this study, the principal calculations for each model, estimation procedure, degree of autocorrelation of the error term and each sample size are presented below.

From Table 1, as the autocorrelation level varies the value of  $\beta's$  becomes unstable as the sample size varies. For the models considered, the  $\beta's$  exhibit damped oscillation in nature with  $\beta_1$  and  $\beta_2$  been negative at n= 20 and 30 respectively.

	$\beta_0$	$eta_1$	$\beta_2$	
Actual	10.0	5.0	- 2.0	
Model 1 n=20	9.9916	4.2268	-2.1552	
n=30	10.3243	-0.6252	0.7172	
n=50	9.6816	4.0752	-0.6024	
n=80	10.5673	0.1606	1.5598	
Model 2 n=20	9.7138	3.8262	-1.8792	
n=30	10.2175	-0.3775	0.7548	
n=50	9.5298	3.0180	0.6547	
n=80	9.9406	0.6641	1.9015	

Table 1. Comparison of the models with the sample sizes

From table 2, when the sample size is 20, the  $\beta_0$  and  $\beta_1$  exhibit a steady downwards positive bias at which model 1 is best, for  $\beta_2$ , it rose steadily with model 1 as best. When the sample size is 30,  $\beta_0$ performed well in the two models with model 5 giving a least bias, in  $\beta_1$ , there is a fluctuation with all values been negative, and at  $\beta_2$ , it rose steadily with model 1 as best. At n=50,  $\beta_0$  and  $\beta_1$  decreased steadily with model one as best and  $\beta_2$  fluctuate as model 1 as best. At n=80,  $\beta_0$  decreased steadily with model 1 as best,  $\beta_1$  and  $\beta_2$  produces a steady increase with all values been positive.

	$\beta_0$	$\beta_1$	$\beta_2$
Actual	10.0	5.0	- 2.0
Model 1 n=20	9.9916	4.2268	-2.1552
Model 2 n=20	9.7138	3.8262	-1.8792
Model 1 n=30	10.3243	-0.6252	0.7172
Model 2 n=30	10.2175	-0.3775	0.7548
Model 1 n=50	9.6816	4.0752	-0.6024
Model 2 n=50	9.5298	3.0180	0.6547
Model 1 n=80	10.5673	0.1606	1.5598
Model 2 n=80	9.9406	0.6641	1.9015

Table 2. Comparison of coefficients of the models based on the same sample size

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From table 3, most of the biases produced are negative for both  $\beta_0$  and  $\beta_1$  while for  $\beta_2$ , the bias increases as the sample sizes increase for each model.

From table 4, model 1 exhibit damped oscillation with n=20 as best while at model 2 there was a sudden decrease after n=30 with n=20 as best.

	$\boldsymbol{\beta}_0$	$\beta_1$	$\beta_2$
Actual	10.0	5.0	- 2.0
Model n=20	-0.0084	-0.7732	-0.1552
n=30	0.3243	-5.6252	2.7172
n=50	-0.3184	-0.0752	1.3976
n=80	0.5673	4.8394	3.5598
Model 2 n=20	-0.2862	-1.1738	0.1208
n=30	0.2175	4.6225	2.7548
n=50	-0.4702	-1.9820	2.6547
n=80	-0.0594	-4.3359	3.9015

**Table 3.** Bias table based on 3.1

Table 4. The MSE and its ranking in ascending order

	$\beta_0$	$\beta_1$	$\beta_2$	Mean
Actual	10.0	5.0	- 2.0	
Model 1 n=20	1	2	1	1.3
n=30	3	4	3	3.3
n=50	2	1	2	1.7
n=80	4	3	4	3
Model 2 n=20	3	1	1	1.7
n=30	2	4	4	3.3
n=50	4	2	2	2.7
n=80	1	3	3	2.3

# 4. Conclusions

In this study, we have examined the performances of the estimators in estimating effect of specification error due to autocorrelation in regression model. Criteria considered are bias, variance and mean square error. Based on the criteria considered, the estimate has been seriously affected by the

autocorrelation as the effects on the models considered are not stable with model 2preferable based on the minimum bias and the MSE at large. The effect of increasing autocorrelation of the error terms on the model is significant i.e as the autocorrelation level increases the model get better. As the sample size increases and the coefficient changes, the effect becomes unstable. It produces an unreliable and less precise estimates i.e. bias and inefficient estimates and the change in the coefficient of the autocorrelation level from 0.5 *and* 0.9 has seriously affected the instability of the results obtained.

Also, our simulation results suggest that the performance of the estimator (OLS) depends on the number of replications. In this research study, we observed a number of unexpected results since some of our findings do not follow a conclusive pattern which reveal that the search for best estimators of models plagues by autocorrelation disturbances could be hazardous.

Further study is being carried out to increase the number of models considered, bootstrap replication, sample size to see if the effect of the autocorrelation could be more noticeable.

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