



## **Association of treatment satisfaction and psychopathological sub-syndromes among involuntary patients with psychotic disorders**

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2 **Association of treatment satisfaction and psychopathological**  
3 **sub-syndromes among involuntary patients with psychotic**  
4 **disorders**

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9 **Abstract**

10 *Purpose* Previous research has shown a link between  
11 treatment satisfaction and global psychopathology in dif-  
12 ferent groups of psychiatric patients. However, neither the  
13 relationship between treatment satisfaction and the sub-  
14 syndromes of global psychopathology nor their temporal  
15 ordering have been explored.

16 *Methods* Participants admitted involuntarily to psychiat-  
17 ric wards in the UK and diagnosed with psychotic disorders  
18 ( $N = 232$ ) were included. Treatment satisfaction and psy-  
19 chopathological sub-syndromes (i.e., manic excitement,  
20 anxiety-depression, negative symptoms, positive symp-  
21 toms) were measured within 1 week and at 1 month after  
22 admission.

23 *Results* Repeated measures ANOVAs showed that higher  
24 treatment satisfaction is associated with lower scores on the  
25 manic excitement, anxiety-depression and positive symp-  
26 toms sub-syndromes, while no significant association was  
27 found for negative symptoms. However, cross-lagged panel  
28 analyses showed that treatment satisfaction predicted  
29 change only in positive symptoms while none of the paths  
30 from the relevant sub-syndromes to treatment satisfaction  
31 was significant.

32 *Conclusion* Treatment satisfaction can be regarded as  
33 an antecedent of changes in positive symptoms only.

These results underline the importance of examining psy- 34  
chopathological sub-syndromes separately as they may 35  
relate differentially to other important correlates of 36  
psychoses. 37

**Keywords** Treatment satisfaction · 39  
BPRS sub-syndromes · Psychoses 40

**Introduction** 41

Treatment satisfaction refers to patients' perceptions con- 42  
cerning their satisfaction and appropriateness of their 43  
treatment [25]. Satisfaction with treatment is critical to 44  
treatment adherence [9] and among the most widely 45  
explored patient-reported outcomes [18]. A link between 46  
treatment satisfaction, assessed within a maximum of 47  
3 days after admission and global psychopathology is 48  
clearly established with higher satisfaction associated with 49  
more favourable outcomes [6, 24, 25]. Involuntary legal 50  
status has consistently been identified as a predictor of 51  
lower satisfaction [10] when compared to patients with 52  
voluntary admission status and among involuntary patients 53  
perceived coercion has been identified as an antecedent of 54  
treatment satisfaction [14]. Thus, the targeting of involun- 55  
tary patients' satisfaction is of clinical relevance, but also an 56  
important ethical issue, as these patients cannot discontinue 57  
their treatment even when they are displeased with it [14]. 58

Does treatment satisfaction influence symptom change 59  
or does symptom change influence treatment satisfaction or 60  
both? Theoretically, it is usually assumed that higher 61  
treatment satisfaction is linked to more symptom 62  
improvement while lower satisfaction is linked to no 63  
improvement or even a deterioration of symptoms. For 64  
example, research has shown higher treatment satisfaction 65

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66 to predict lower global psychopathology among patients  
 67 with a range of psychiatric diagnoses [6, 24, 25]. None-  
 68 theless, satisfaction has also been modelled as an outcome.  
 69 For example, Katsakou et al. [14] showed that patients who  
 70 perceived less coercion at admission and during hospital  
 71 treatment and patients with more symptom improvement  
 72 expressed higher treatment satisfaction, see Bjorngaard  
 73 et al. [4] and Shiva et al. [33]. As described by Burkholder  
 74 and Harlow [7] structural equation models that examine  
 75 cross-lagged time paths between variables can help to  
 76 determine their temporal ordering. If cross-lagged paths  
 77 from both treatment satisfaction to symptoms and from  
 78 symptoms to treatment satisfaction were statistically sig-  
 79 nificant, a reciprocal association between the constructs  
 80 would be suggested. However, if only the path from  
 81 treatment satisfaction to symptoms is statistically signifi-  
 82 cant, it may be concluded that treatment satisfaction pre-  
 83 ceedes symptoms and not the other way round. Conversely,  
 84 if only the path from symptoms to treatment satisfaction is  
 85 significant, symptoms may be seen as an antecedent of  
 86 treatment satisfaction.

87 These studies exploring treatment satisfaction have used  
 88 global psychopathology as the criterion. However, among  
 89 patients with psychotic disorders, global measures of psy-  
 90 chopathology comprise at least four interpretable sub-  
 91 syndromes, namely manic excitement, anxiety-depression,  
 92 negative symptoms and positive symptoms, which may be  
 93 influenced by separate processes and aetiologies [30].  
 94 Citing Lachar et al. [15] and Van der Does et al. [34],  
 95 Shafer [32] advocates examination of these sub-syndromes  
 96 independently, arguing that using global scores may mask  
 97 important treatment effects and specific areas of symptom  
 98 change. Nonetheless, due to the dearth of research on sub-  
 99 syndromes, specific mechanisms and therefore hypotheses  
 100 for each sub-syndrome cannot be specified. However, in  
 101 general, admission onto a psychiatric ward is expected to  
 102 promote clinical improvement including symptom out-  
 103 comes [3].

#### 104 Aims and hypotheses

105 The relationship between subjective treatment satisfaction  
 106 and the facets of global psychopathology have not been  
 107 explored. Moreover, the direction of the association  
 108 between treatment satisfaction and psychopathological  
 109 symptoms has not been tested. A longitudinal design,  
 110 where both treatment satisfaction and psychopathological  
 111 sub-syndromes are measured repeatedly across time can  
 112 facilitate exploration of these questions.

113 Following on from this, three hypotheses were tested:

114 i) Patients with higher treatment satisfaction during the  
 115 first week of admission will report lower scores on the

116 manic excitement, anxiety-depression, negative and  
 117 positive sub-syndromes overall (i.e., between the first  
 118 week of admission and 1 month post admission) than  
 119 those with lower treatment satisfaction.

120 ii) In a cross-lagged panel design with latent variables,  
 121 higher treatment satisfaction will predict symptom  
 122 improvement between the first week of admission and  
 123 1 month post-admission.

124 iii) In a cross-lagged panel design with latent variables,  
 125 fewer symptoms will predict higher satisfaction with  
 126 treatment between the first week of admission and  
 127 1 month post-admission.

## 128 Method

### 129 Participants

130 All potential participants had been admitted involuntarily  
 131 to a psychiatric ward in the UK between July 2003 and July  
 132 2005 and were recruited for a larger study for which  
 133 detailed inclusion criteria and recruitment process have  
 134 been described elsewhere [26]. Data collection for the  
 135 initial study was approved by the multicentre research  
 136 ethics committee and all participants gave written informed  
 137 consent to take part. Compared to all eligible patients,  
 138 participants interviewed at baseline were more likely to be  
 139 younger and more likely to be male [cf., 26]. Of the 778  
 140 patients interviewed at baseline, only those diagnosed with  
 141 schizophrenia or other psychosis, according to the ICD-10  
 142 categories (i.e., F20-29) were included ( $N = 383$ ). A mean  
 143 age of 35.91 ( $\pm 10.94$ ) was reported and 276 (72%) of the  
 144 participants were male.

### 145 Measures

146 In baseline interviews, participants were asked to provide  
 147 socio-demographic information including ethnicity (the  
 148 United Kingdom census 2001 categories collapsed into 2  
 149 categories: white versus ethnic minority), and education (4  
 150 categories: no qualification, GCSE grades A–C, 'A' level  
 151 or equivalent, and degree). Information on the total length  
 152 of stay (in days) was also collected from medical records.  
 153 Measures of treatment satisfaction and psychopathological  
 154 sub-syndromes were each measured within 1 week and at  
 155 1 month post-admission. For each construct, multi-item  
 156 scale scores were computed by averaging participants'  
 157 responses across the relevant items.

158 The Client's Assessment of Treatment Scale (CAT) was  
 159 used to assess treatment satisfaction [12, 25] which has  
 160 been used in studies with psychiatric inpatients. The scale  
 161 assesses patients' subjective satisfaction and perceptions of

162 appropriateness of their treatment using 7 items (e.g., “Do  
163 you believe you are receiving the right treatment/care for  
164 you here?”, “Are relations with other staff members  
165 pleasant for you?”, “Does your psychiatrist understand you  
166 and is he/she engaged in your treatment/care?”). Each item  
167 is rated on a 11-point Lickert-type scale that ranges from 0  
168 ‘not at all’ to 10 ‘yes entirely’ ( $M = 5.51 \pm 2.77$  and  
169  $6.05 \pm 2.61$  at week 1 and 1 month, respectively).

170 Psychopathological symptoms were researcher rated  
171 using the 24-item Brief Psychiatric Rating Scale (BPRS)  
172 [35]. Items assess symptom severity on 7-point Likert-type  
173 scales with end points that range from ‘not present’ to  
174 ‘extremely severe’. Sub-syndromes were indexed using  
175 a factor analytic solution of the BPRS among patients  
176 with schizophrenia living in the UK [24]. Manic excite-  
177 ment ( $M = 2.10 \pm 0.61$  and  $1.56 \pm 0.56$  at week 1 and  
178 1 month, respectively) was assessed by 9 items (e.g.,  
179 hostility, elevated mood), anxiety/depression ( $M = 2.28 \pm$   
180  $0.92$  and  $2.03 \pm 0.88$  at week 1 and 1 month, respectively)  
181 by 6 items (e.g., somatic concern, anxiety), negative  
182 symptoms ( $M = 1.79 \pm 0.88$  and  $1.63 \pm 0.72$  at week 1  
183 and 1 month, respectively) by 4 items (e.g., disorientation,  
184 blunted affect) and positive symptoms ( $M = 3.18 \pm 1.22$   
185 and  $2.20 \pm 1.22$  at week 1 and 1 month, respectively)  
186 using 5 items (e.g., grandiosity, suspiciousness).

## 187 Analytic strategy

188 Prior to testing the study hypotheses, listwise deletion  
189 procedures were used to account for missing data. Thus, in  
190 order to assess the representative of our samples  $t$  test and  
191  $\chi^2$  analyses were conducted to compare those eligible for  
192 the study ( $N = 383$ ) and participants for whom complete  
193 data were available at both points of time ( $N = 232$ ).

194 Following Luszczynska et al. [16] the hypotheses were  
195 tested in 3 analytic steps. First, correlations between the  
196 variables were examined. Second, repeated measures  
197 analyses of variance across two time points with treatment  
198 satisfaction as a between subjects factor (two levels) was  
199 used to examine the association between initial treatment  
200 satisfaction and each sub-syndrome over time (hypothesis  
201 1). Third, a two-step structural equation model (SEM) [1]  
202 was used to assess the temporal ordering of treatment  
203 satisfaction and each sub-syndrome (hypotheses 2 and 3).

204 The EQS 6 programme [3] was used to test the temporal  
205 ordering of treatment satisfaction and each sub-syndrome  
206 using the maximum likelihood method for all analyses. A  
207 two wave cross-lagged panel model with a 3-week time lag  
208 was estimated. A two-step approach to SEM was used to  
209 assess the validity and reliability of the constructs before  
210 their use in the structural model [1].

211 In the measurement models, treatment satisfaction and  
212 the focal sub-syndrome at both time points were modelled

simultaneously in a single model. Ideally, parameter 213  
loadings for each separate item on the corresponding latent 214  
factors would be estimated. However, the size of the 215  
sample was too small for the number of estimated param- 216  
eters that such a model would produce, so an item par- 217  
celling strategy [2] was adopted. Specifically, we created 218  
three indicators for measures of treatment satisfaction and 219  
each psychopathological sub-syndrome (at each time point) 220  
using randomly selected item parcels. The same items were 221  
included in the parcels at each time point (to ensure that the 222  
nature of the constructs did not change over time). Refer- 223  
ence indicators for each latent variable were created by 224  
fixing the highest indicator’s loading to 1 and as is the 225  
usual case in confirmatory factor analysis, the latent con- 226  
structs were allowed to co-vary. Error terms across time 227  
points for the same indicator were allowed to co-vary, 228  
where the Lagrange multiplier test indicated that this would 229  
lead to a statistically significant improvement in model fit 230  
[21]. 231

Subsequent path models examined crossover paths 232  
between satisfaction and the focal sub-syndrome. Specifi- 233  
cally, the association between treatment satisfaction during 234  
week 1 of admission (time 1) and the focal sub-syndrome 235  
at 1 month post-admission (time 2) was compared to the 236  
relevant association between the focal sub-syndrome dur- 237  
ing week 1 of admission (time 1) and treatment satisfaction 238  
measured at 1 month post-admission (time 2). Auto- 239  
regression coefficients were also specified to control for 240  
covariance stability between the same constructs over time. 241

As the chi-square goodness of fit statistic is sensitive to 242  
sample size [17] additional recommended indices for 243  
goodness of fit and cutoffs [11] were used to evaluate the 244  
adequacy of the models. Specifically, in addition to the  $\chi^2$  245  
test statistic, the comparative fit index (CFI), non-normed 246  
fit index (NNFI) and the root-mean square error of 247  
approximation (RMSEA) are reported. A non-significant  $\chi^2$  248  
value ( $p > 0.05$ ), CFI and NNFI values of 0.90 (or above) 249  
and a RMSEA of 0.08 (or lower) reflect adequate model fit. 250

## 251 Results

Comparison between participants eligible for the study 252  
( $N = 383$ ) and participants for whom complete data were 253  
available at both time points ( $N = 232$ ) showed that these 254  
groups differed neither in gender, age, ethnicity, education 255  
and length of stay. Table 1 shows the corresponding 256  
descriptive statistics and frequencies. 257

Table 2 presents the correlations between the study 258  
variables. Higher treatment satisfaction measured during 259  
week 1 was associated with lower global psychopathology, 260  
manic excitement, anxiety-depression and positive symp- 261  
toms at both time points ( $r$  range from  $-0.12$  to  $-0.19$ ). 262

**Table 1** Descriptive statistics comparing patients at baseline and those with complete data at 1 month

	Baseline <i>N</i> = 383	1 month <i>N</i> = 232
Male		
<i>N</i> (%)	72	72
Age on admission		
Mean (SD)	35.91 (10.95)	35.84 (11.47)
Ethnicity		
White (%)	63	63
Education		
No qualifications (%)	31	30
A–C GCSEs (%)	24	25
'A' level or equivalent (%)	36	36
Degree (%)	9	10
Length of stay		
Mean (SD)	88.91 (84.09)	94.72 (85.42)

263 Contrary to expectation, lower treatment satisfaction during  
264 week 1 was associated with fewer negative symptoms  
265 although this did not reach a conventional level of statisti-  
266 cal significance. Neither global psychopathology nor the  
267 four sub-syndromes measured within week 1 were associ-  
268 ated with treatment satisfaction at 1 month (*r* range from  
269 –0.06 to 0.05). These results support our first hypothesis  
270 for manic excitement, anxiety-depression and positive sub-

syndromes and indicate that treatment satisfaction precedes  
symptoms (hypotheses 2) rather than the reverse temporal  
hypothesis (hypothesis 3).

Changes in symptoms over time depending  
on treatment satisfaction

Repeated measures analyses of variance across two time  
points were conducted for each of the four sub-syndromes  
with treatment satisfaction (measured during week 1) as a  
between-subjects factor (two levels). For this analyses,  
treatment satisfaction scores were standardised and partic-  
ipants scoring above (*N* = 113) and below zero  
(*N* = 119), respectively, were categorised into high- and  
low-satisfaction groups.

The mean score for each sub-syndrome at high and low  
levels of satisfaction are shown in Fig. 1. Each sub-syn-  
drome changed statistically significantly over time, reduc-  
ing from week 1 to 1 month post-admission,  $F(1, 230) = 149.80$  ( $r = 0.63$ ),  
 $22.69$  ( $r = 0.30$ ),  $10.67$  ( $r = 0.21$ ) and  $129.21$  ( $r = 0.60$ ) (all  $p < 0.01$ ), respec-  
tively, for manic excitement, anxiety-depression, negative  
symptoms and positive symptoms. With exception of nega-  
tive symptoms,  $F(1, 230) = 2.52$ ,  $p > 0.05$  ( $r = 0.10$ ),  
patients with higher treatment satisfaction reported fewer  
symptoms overall (i.e., across both time points),  $F(1, 230) = 4.74$ ,  
( $r = 0.14$ ),  $3.99$  ( $r = 0.13$ ) and  $5.59$

**Table 2** Correlations among the study variables during week 1 (T1) and 1 month (T2) post involuntary admission for patients with psychoses

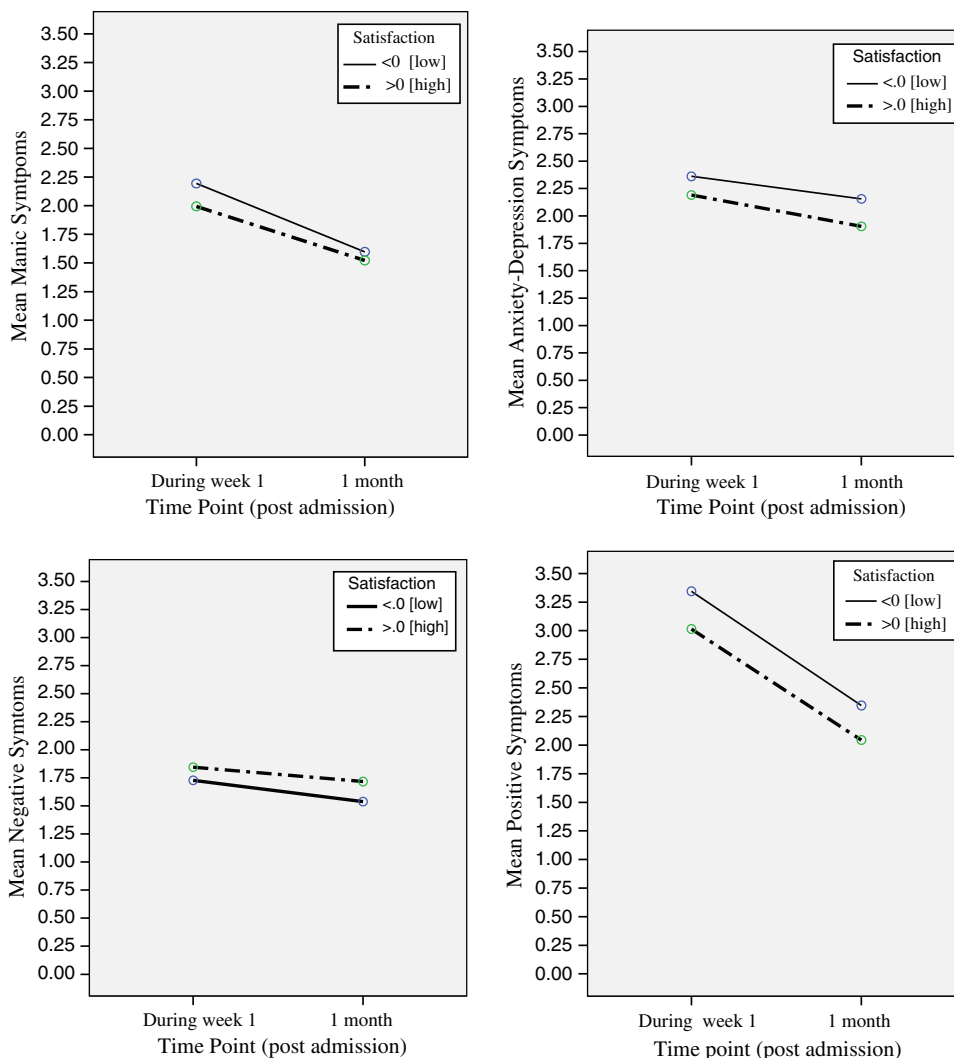
	T-sat T1	T-sat T2	BPRS T1	BPRS T2	Manic T1	Manic T2	Anx-dep T1	Anx-dep T2	Negative T1	Negative T2	Positive T1	Positive T2
T-sat T1		0.52**	-0.19**	-0.15*	-0.18**	-0.12***	-0.15*	-0.14*	0.09	0.11 <sup>†</sup>	-0.15*	-0.17*
T-sat T2			-0.02	-0.29**	0.02	-0.23**	-0.06	-0.20**	0.05	-0.04	-0.05	0.25**
BPRS T1				0.51**	0.69**	0.29**	0.55**	0.38**	0.40**	0.25**	0.73**	0.42**
BPRS T2					0.22**	0.71**	0.41**	0.68**	0.31**	0.42**	0.31**	0.83**
Manic T1						0.36**	0.08	-0.02	0.11	0.02	0.40**	0.19**
Manic T2							0.14*	0.22**	0.12	0.12	0.09	0.49**
Anx-dep T1								0.62**	0.07	0.10	0.12	0.20**
Anx-dep T2									0.11	0.17*	0.18*	0.36**
Negative T1										0.58**	0.10	0.20**
Negative T2											0.08	0.22**
Positive T1												0.42**
Positive T2												

T1 time 1, T2 time 2, T-Sat treatment satisfaction, BPRS brief psychiatric rating scale (mean score), Manic manic excitement, Anx-dep anxiety depression

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$ , <sup>†</sup>  $p < 0.09$ , *N* = 232



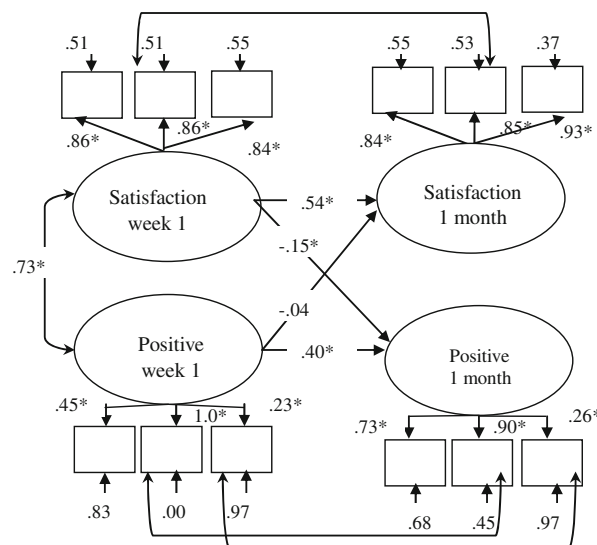
**Fig. 1** Mean scores on the sub-syndromes during week 1 and at 1 month post-admission with initial treatment satisfaction as the between-subjects factor



296 ( $r = 0.16$ ) for manic excitement, anxiety-depression and  
 297 positive symptoms, respectively. For each sub-syndrome  
 298 the interaction between time and satisfaction was insignif-  
 299 icant indicating that the reduction in symptoms was similar  
 300 for participants high or low in satisfaction,  $F(1,$   
 301  $230) = 2.05, 0.62, 0.40$  and  $0.30$  (all  $p > 0.05$ ) for manic  
 302 excitement, anxiety-depression, negative symptoms and  
 303 positive symptoms, respectively. Thus, hypothesis 1 was  
 304 supported for manic excitement, anxiety-depression and  
 305 positive symptoms only. As the negative sub-syndrome was  
 306 unrelated to treatment satisfaction it was excluded from all  
 307 further analyses (Fig. 2).

308 Table 3 shows that the measurement model for each  
 309 sub-syndrome fitted the data reasonably well. The factor  
 310 loading of each indicator to its hypothesised latent factor  
 311 was significant providing evidence of a stable structure in  
 312 each group.

313 Table 4 presents the goodness of fit indices for the cross-  
 314 lagged panel models. Results show that for manic excite-  
 315 ment, neither of the cross-lagged effects was statistically



**Fig. 2** Standardised parameter estimates for the full SEM model of treatment satisfaction and positive symptoms among patients with psychotic disorders

Author Proof

**Table 3**  $\chi^2$  and fit indices for the measurement models

Model	$\chi^2$	df, N	RMSEA	90% CI for RMSEA	NNFI	CFI
Manic excitement	63.213, $p = 0.05$	46, 232	0.04	0.00–0.06	0.98	0.98
Anxiety-depression	73.147, $p = 0.00$	44, 232	0.05	0.03–0.08	0.97	0.98
Positive symptoms	53.52, $p = 0.18$	45, 232	0.03	0.00–0.06	0.99	0.99

RMSEA root-mean-square error of approximation, CI confidence interval; NNFI non-normed fit index, CFI comparative fit index

**Table 4**  $\chi^2$  and fit indices for the cross-lagged panel models

Model	$\chi^2$	df, N	RMSEA	90% CI for RMSEA	NNFI	CFI
Manic excitement	80.044, $p = 0.00$	47, 232	0.06	0.03–0.08	0.96	0.97
Anxiety-depression	89.006, $p = 0.00$	46, 232	0.04	0.04–0.08	0.95	0.97
Positive symptoms	62.125, $p = 0.06$	46, 232	0.04	0.00–0.06	0.98	0.99

RMSEA root-mean-square error of approximation, CI confidence interval, NNFI non-normed fit index, CFI comparative fit index

316 significant ( $\beta = 0.11$  and  $0.02$  for the path from satisfaction  
317 during week 1 to manic excitement at 1 month and from  
318 manic excitement during week 1 to treatment satisfaction at  
319 1 month, respectively). A similar pattern of results was  
320 observed for the anxiety-depression sub-syndrome. Spe-  
321 cifically, neither the path from treatment satisfaction during  
322 week 1 to anxiety-depression at 1 month ( $\beta = -0.04$ ) or  
323 anxiety-depression during week 1 to treatment satisfaction  
324 at 1 month ( $\beta = -0.04$ ) were statistically significant. In  
325 contrast, results for the positive sub-syndrome revealed a  
326 statistically significant negative beta coefficient for the path  
327 from treatment satisfaction during week 1 to positive  
328 symptoms at 1 month ( $\beta = -0.15$ ) while the path for the  
329 reverse temporal ordering was negligible and insignificant  
330 both in size and in statistically ( $\beta = -0.04$ ). Supporting  
331 this, the model fit indices reported in Table 4 show that the  
332 cross-lagged model for the positive sub-syndrome fit the  
333 data well,  $\chi^2(46, 232) = 62.13$ ,  $p = 0.06$ , CFI = 0.99,  
334 NNFI = 0.98, RMSEA = 0.04 (90% CI 0.00–0.06). Thus,  
335 hypothesis 2 was supported for the positive sub-syndrome  
336 only while no support for hypothesis 3 was found.

## 337 Discussion

338 The relationship between treatment satisfaction and psy-  
339 chopathological sub-syndromes were examined among  
340 involuntary in-patients in the UK with psychotic disorders.  
341 With exception of the negative sub-syndrome, participants  
342 reporting higher treatment satisfaction exhibited fewer  
343 symptoms compared to those with lower treatment satis-  
344 faction. Thus, the first hypothesis was supported for all sub-  
345 syndromes except negative symptoms. The cross-lagged  
346 panel analysis showed that treatment satisfaction predicted  
347 change in only the positive symptom sub-syndrome pro-  
348 viding support for our 2nd hypothesis. The reverse

temporal hypothesis (hypothesis 3) was not supported for 349  
any of the sub-syndromes. 350

The finding that treatment satisfaction relates differen- 351  
tially to the sub-syndromes of psychoses is new and adds to 352  
an increasing body of research emphasising the importance 353  
of examining sub-syndromes separately [33]. Indeed, dis- 354  
regarding different symptom dimensions may mask iso- 355  
lated areas of symptom change and dilute the global effect. 356  
A second new finding is that change in treatment satis- 357  
faction predicts change in scores on the positive sub-syn- 358  
drome, confirming the assumption that satisfaction can be 359  
regarded as an antecedent of positive symptoms. However, 360  
neither of the cross-lagged paths was significant for manic 361  
excitement and anxiety-depression sub-syndromes, sug- 362  
gesting spurious time-lagged correlations arising from 363  
significant concurrent associations and the stability of these 364  
constructs over time. 365

The reasons why treatment satisfaction should influence 366  
change in positive symptoms versus the other sub-syn- 367  
dromes is unclear. However, the finding is consistent with 368  
previous research showing that positive symptoms may be 369  
more malleable and amenable to intervention [19]. In the 370  
current study sample, only 19% ( $N = 44$ ) left hospital prior 371  
to the assessment at 1 month and of these, the mean length 372  
of stay was 19.18 days ( $SD = 7.04$ ). Consequently, as 373  
medication adherence was involuntary and regulated 374  
among the majority of patients' adherence is an unlikely 375  
mediator. 376

While one can only speculate about the mechanisms of 377  
change we believe that individual difference and social 378  
factors may also play an important role in the relationship 379  
between treatment satisfaction and symptom improvement. 380  
For example, patients' perceptions of autonomy may 381  
mediate the relationship between treatment satisfaction and 382  
improvement in positive symptoms. Indeed, according to 383  
some theories, (e.g., self determination theory) [31] 384

385 autonomy supportive environments (e.g., high levels of  
386 perceived control among patients) have been shown to  
387 facilitate motivation for treatment [5]. This speculation  
388 certainly coincides with the finding that higher perceived  
389 coercion among involuntary patients is linked to lower  
390 levels of satisfaction with treatment [14]. Supporting this,  
391 in a recent qualitative study objectification and marginali-  
392 sation of the patient was identified by patients as one of key  
393 themes concerning their care [28].

394 It is noteworthy that the CAT scale includes components  
395 of therapeutic alliance in addition to more general aspects of  
396 treatment satisfaction and it might thus be argued that these  
397 constructs are synonymous. Indeed therapeutic alliance has  
398 been shown to explain similar proportions of symptom  
399 improvement to that found in the current study [8]. None-  
400 theless, recent research has shown that although therapeutic  
401 alliance and treatment satisfaction share common variance,  
402 they too provide distinct information from this overlap [29].  
403 However, incremental predictive validity studies including  
404 both of these constructs in addition to other predictors of  
405 symptom reduction, such as unmet needs for care [22] and  
406 subjective quality of life [20] among psychiatric patients are  
407 not widely reported thus more research is needed to ascer-  
408 tain their relative importance.

409 The research reported here adds to the growing body of  
410 evidence indicating that subjective patient reports are  
411 predictive of important clinical outcomes. As treatment  
412 satisfaction is relatively easy to elicit and could be added  
413 easily to routine clinical practice these findings may have  
414 considerable practical application. Nonetheless, the effect  
415 size estimate between treatment satisfaction and improve-  
416 ment in positive symptoms was relatively small ( $\beta =$   
417  $-0.15$ ). In any case the findings indicate that it may be  
418 worth developing interventions to enhance treatment sat-  
419 isfaction. If such interventions could be developed and  
420 were found to be effective, they might also shed further  
421 theoretical light on the psychological antecedents of posi-  
422 tive symptoms, e.g., by identifying moderators and medi-  
423 ators of treatment re-training effectiveness. This could also  
424 facilitate assessment of the potential risks, gains and cost  
425 effectiveness of such interventions and therefore assess-  
426 ment of their practical utility.

427 The use of latent variable SEM allowed examination of  
428 relationships between constructs with measures that were  
429 relatively free of measurement error. Additionally, the  
430 sample size was large and comprised a relatively large and  
431 diagnostically homogeneous sample. Moreover, the use of  
432 researcher-rated rather than self-reported outcomes reduced  
433 the likelihood of artificially inflating effect size estimates  
434 resulting from common method variance. Nonetheless,  
435 data were only available for those patients willing to take  
436 part in academic research which may have introduced a  
437 selection bias. Also, although the sample size is impressive

438 for this particular group of patients the large number  
439 required for statistical modelling meant that examination of  
440 each item individually to its respective factor was not  
441 feasible in the SEM. This is important, as while the CAT  
442 items have good face validity and predictive utility, in  
443 addition to high internal consistency reliability [26], the  
444 factorial validity of the CAT is yet to be established.

445 To test the generalisability of the study findings, future  
446 research is needed to replicate the current findings in  
447 samples with different diagnoses and for patients in dif-  
448 ferent treatment settings. Moreover, theory-based research  
449 could help to locate the specific mechanisms that lead to  
450 change in positive symptoms. For these purposes, future  
451 studies could test psychological theories (such as self-  
452 determination theory) which provide a theoretical frame-  
453 work for exploring these relationships.

454 Understandably, clinicians might think that immediate  
455 patient satisfaction is not that relevant among patients  
456 compulsory admitted. However, this study emphasises that  
457 what patients think about their care within the first week of  
458 treatment is an indicator of changes in positive symptoms  
459 at 1 month post-admission and thus, could be considered  
460 even when symptom levels are often still high and the  
461 situation tense.

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470 **Conflict of interest statement** All authors declare that they have no  
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