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Keywords: breast cancer; survival; aspirin; general practice; record linkage

Aspirin use and survival after the diagnosis of breast cancer: a population-based cohort study

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Background: Aspirin use has been associated with a reduced cancer incidence and fewer deaths from cancer. This study examined whether women with breast cancer prescribed aspirin postdiagnosis had improved survival.

Methods: An observational, population cohort study was undertaken using data linkage of cancer registry, dispensed prescriptions and death records in Tayside, Scotland. All community prescriptions for aspirin in women with breast cancer were extracted and use postdiagnosis for each individual examined using Cox's proportional hazard models. The main outcome measures were all-cause mortality and breast cancer-specific mortality.

Results: Four thousand six hundred and twenty-seven patients diagnosed with breast cancer between 1 January 1998 and 31 December 2008 were followed up until 28 February 2010. Median age at diagnosis was 62 (IQR 52–74). One thousand eight hundred and two (39%) deaths were recorded, with 815 (18%) attributed to breast cancer. One thousand and thirty-five (22%) patients were prescribed aspirin postdiagnosis. Such aspirin use was associated with lower risk of all-cause mortality (HR = 0.53, 95% CI = 0.45–0.63, $P < 0.001$) and breast cancer-specific mortality (HR = 0.42, 95% CI = 0.31–0.55, $P < 0.001$) after adjusting for age, socioeconomic status, TNM stage, tumour grade, oestrogen receptor status, surgery, radiotherapy, chemotherapy, adjuvant endocrine therapy and aspirin use prediagnosis.

Conclusions: Aspirin use postdiagnosis of breast cancer may reduce both all-cause and breast cancer-specific mortality. Further investigation seeking a causal relationship and which subgroups of patients benefit most await ongoing randomised controlled trials.

Breast cancer is the most common malignancy in the United Kingdom, with over 40 000 women diagnosed in 2008 (Maddams *et al*, 2009). The molecular events leading to the initiation and progression of breast cancer are not completely understood; however, tissue enzymes such as aromatases and prostaglandins may have a role in the development of the disease. (Kulendran *et al*, 2009; Hoellen *et al*, 2011)

Aspirin is a common non-steroidal anti-inflammatory drug (NSAID) often used for analgesia or at low dosage as an antiplatelet agent for the prevention of myocardial infarction and stroke.

Aspirin acts by irreversibly inhibiting cyclooxygenase (PTGS, previously COX-1) and modifies the activity of PTGS-2 required for the synthesis of prostaglandins. Other NSAIDs, such as ibuprofen, reversibly inhibit cyclooxygenase. Non-steroidal anti-inflammatory drugs, particularly aspirin, may have chemopreventive or even therapeutic properties for several common types of cancer (Gupta and DuBois, 2001; Rostom *et al*, 2007; Takkouche *et al*, 2008).

Several case-control studies have shown a significant reduction in the risk of breast cancer with NSAID and aspirin use (Harris *et al*, 2006; Kirsh *et al*, 2007). However, prospective

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studies have given mixed results; aspirin has been shown to have no association (Gill *et al*, 2007; Jacobs *et al*, 2007), a decreased risk (Schreinemachers and Everson, 1994; Harris *et al*, 2003) or an increased risk (Friis *et al*, 2008); non-aspirin NSAIDs have been associated with a reduced risk (Schreinemachers and Everson, 1994; Harris *et al*, 2003; Gill *et al*, 2007).

Aspirin, but not NSAIDs, may interact with other cellular processes, independent of aromatase and oestrogen expression (Bardia *et al*, 2011). An indirect effect of aspirin on mammary serpin (maspin) has been implicated in the prevention and control of breast cancer (Zou *et al*, 1994) and in an animal model by restoring nitrous oxide synthesis, increasing maspin production (Bhattacharyya *et al*, 2010). Overexpression of PTGS-2 in tumours has also been linked to promoting angiogenesis and inhibiting apoptosis (Sheng *et al*, 1998; Rozic *et al*, 2001), further implicating that aspirin could prevent growth, metastasis and recurrence.

In the therapeutic setting, a significant reduction in relative risk by taking aspirin for 2 or more days per week was demonstrated both for breast cancer metastasis and breast cancer mortality (Holmes *et al*, 2010). In a questionnaire-based cohort study of 41 000 postmenopausal women (aged 55–69 years), aspirin was associated with a reduction in all-cause mortality and breast cancer mortality (Blair *et al*, 2007).

The aim of this study was to examine if aspirin use in a population-based cohort of women with breast cancer postdiagnosis was associated with all-cause mortality or breast cancer-specific mortality.

MATERIALS AND METHODS

The Health Informatics Centre (HIC) at the University of Dundee holds health-related databases on all 400 000 residents of the Tayside region, Scotland. All women in Tayside diagnosed with primary invasive breast cancer (ICD10 classifications C50.0–C50.9 or ICD9 classifications 174.0–174.9) between 1 January 1993 and 31 December 2008 and who remained a resident or died in Tayside were identified and studied from the date of diagnosis to either death or the study end date (Makubate *et al*, 2013). Women with a previous diagnosis of any cancer were excluded. Individual patients were identified and linked to encashed prescribing, cancer registry and audit, death certificate and demographic records. Patients were classed as dying from breast cancer if this was listed on the death certificate as the underlying cause of death. From the encashed prescribing data set from 1 January 1993 to 28 February 2010, aspirin prescriptions were identified and use categorised into: never, prediagnosis only, post breast cancer diagnosis only or pre- plus postdiagnosis groups. Aspirin prescribing postdiagnosis included the prescription coverage, calculated from the number of tablets dispensed and the directions for use, to determine the number of days the prescription would last if the patient took the tablets as directed by the prescribing clinician. Thus, a prescription of 56 tablets with directions of one tablet two times daily would have a coverage of 28 days.

Patient follow-up postdiagnosis was split into periods of aspirin use, or non-use, based on aspirin coverage for each individual until death or the end of follow-up. The total duration of aspirin use was calculated from the date of the first prescription postdiagnosis to the end of coverage of the last prescription. Adherence to aspirin was calculated by dividing the total coverage of the aspirin prescriptions by the total duration of its use (Makubate *et al*, 2013).

Ethical approval was granted by the Tayside Committee on Medical Research Ethics and the NHS Tayside Caldicott Guardian according to prior arrangements within the HIC (CA/FB HIC Ethics letter dated 3 February 2010).

Statistical analysis. Data were described as the number of subjects (percentages) for categorical variables and mean with standard

deviation (s.d.) for continuous variables. Where continuous variables did not follow a normal distribution, they were tested using the Shapiro–Wilks test for skewness, and the median and interquartile range were reported. Likelihood ratios and χ^2 tests for trend (χ^2 trend, degrees of freedom (d.f.), probability (*P*)) were reported for differences in distribution of the population with *n*-ordered categories, otherwise Pearson's χ^2 test for differences was used (χ^2 , d.f., *P*).

Cox's proportional hazards models were utilised to estimate hazard ratios (HRs) and 95% confidence intervals (CIs) for each unadjusted and adjusted covariate for cancer diagnosis. Patients were followed up from cancer diagnosis until the date of death or the end of the study. Individual patients who used aspirin postdiagnosis had several linked records each reflecting a period of aspirin use, or no use, which provided continuous follow-up. This marker of aspirin use was included in Cox's proportional hazards model to allow for the effect over time since diagnosis to be accurately examined. The proportional hazards assumption was assessed using trend tests of the Schoenfeld residuals. The multiple regression analysis allowed for age, socioeconomic status (SES), cancer stage and grade at diagnosis, ER status, surgery, radiotherapy, chemotherapy, adjuvant endocrine therapy and aspirin use prediagnosis.

All statistical analyses were performed using Stata version 11 (StataCorp, 2009, Stata Statistical Software: Release 11, College Station TSL, TX, USA).

RESULTS

There were 4627 women diagnosed with incident breast cancer over 11 years, 1 January 1998 to 31 December 2008 inclusive (see Table 1); follow-up was until death or the end of the study period (28 February 2010). The median age at diagnosis was 62 years (IQR 52–74) and patients were followed up postdiagnosis for 31 444 patient years in total, with a median length of follow-up of 5.7 years (IQR 3.0–10.1). One thousand eight hundred and two (39.0%) patients died during the study period, with 815 (17.6%) attributed to breast cancer from death certificate records; 3803 (82%) patients followed up until death or for a minimum of 5 years postdiagnosis.

There were 682 women (14.7%) who took aspirin before diagnosis and 1035 (22.4%) took aspirin postdiagnosis of breast cancer, with a median of 2.4 years use postdiagnosis (IQR 0.7–5.0 years). Calculated adherence to aspirin over the entire period was high, with a median figure of 95.4% (IQR 81.5–100%). The majority, 27 484 (99%), of aspirin prescriptions postdiagnosis were for 75 mg dosage. Other characteristics of patients by aspirin use pre- and postdiagnosis are shown in Table 2. Patients who took aspirin postdiagnosis were older (χ^2 trend = 400.7, d.f. = 5, $P < 0.001$), had higher SES (χ^2 trend = 50.1, d.f. = 5, $P < 0.001$), lower tumour stage ($\chi^2 = 13.4$, d.f. = 4, $P = 0.009$), less metastases ($\chi^2 = 9.8$, d.f. = 2, $P = 0.007$), lower grade tumours ($\chi^2 = 9.5$, d.f. = 3, $P = 0.023$), were less likely to have radiotherapy ($\chi^2 = 27.0$, d.f. = 2, $P < 0.001$) or chemotherapy ($\chi^2 = 121.2$, d.f. = 2, $P < 0.001$) or surgery ($\chi^2 = 26.5$, d.f. = 2, $P < 0.001$) and were more likely to have adjuvant endocrine therapy ($\chi^2 = 74.3$, d.f. = 3, $P < 0.001$) (see Table 2).

Using Cox's regression models, an unadjusted model for all-cause mortality in patients taking aspirin after diagnosis was created along with a multiple regression model adjusted for age, SES, tumour characteristics, surgery, radiotherapy, chemotherapy, adjuvant endocrine therapy and aspirin use prediagnosis. The adjusted model showed that aspirin reduced the risk of all-cause mortality (HR = 0.53, 95% CI = 0.45–0.63, $P < 0.001$) (Table 3). Increasing age, tumour stage, lymph node involvement, metastasis and tumour grade, as well as having chemotherapy were associated

Table 1. Characteristics of patient cohort

	Number of women (%)	All-cause mortality (%)	Breast cancer mortality (%)
Total number of women	4627	1802 (39.0)	815 (15.0)
Age at diagnosis (years)			
<40	207 (4.5)	76 (36.7)	54 (26.1)
40–49	641 (13.9)	156 (24.3)	113 (17.6)
50–59	1100 (23.8)	241 (21.9)	142 (12.9)
60–69	1118 (24.2)	342 (30.6)	151 (13.5)
70–79	883 (19.1)	471 (53.3)	172 (19.5)
80+	675 (14.6)	515 (76.3)	183 (27.1)
Unknown	3 (0.1)	1 (33.3)	0
Socioeconomic status (SCSIMD5)			
1 (most deprived)	599 (12.9)	279 (46.6)	132 (22.0)
2	599 (12.9)	256 (42.7)	115 (19.2)
3	731 (15.8)	299 (40.9)	142 (19.4)
4	1445 (31.2)	577 (39.9)	241 (16.7)
5 (most affluent)	969 (20.9)	339 (35.0)	166 (17.1)
Unknown	284 (6.1)	52 (18.3)	19 (6.7)
Tumour stage			
1	1032 (22.3)	277 (26.8)	103 (10.0)
2	1211 (26.2)	553 (45.7)	269 (22.2)
3	259 (5.6)	154 (59.5)	89 (34.4)
4	322 (6.9)	244 (75.8)	152 (47.2)
Unknown	1803 (39.0)	574 (31.8)	202 (11.2)
Node status			
N0	2595 (56.1)	887 (34.2)	399 (15.4)
N1	511 (11.0)	318 (62.2)	189 (37.0)
N2	100 (2.2)	78 (78.0)	48 (48.0)
Unknown	1421 (30.7)	519 (36.5)	179 (12.6)
Metastases			
No	2959 (64.0)	1122 (37.9)	520 (17.6)
Yes	154 (3.3)	137 (89.0)	89 (57.8)
Unknown	1514 (32.7)	543 (35.9)	206 (13.6)
Tumour grade			
G1	469 (10.1)	113 (24.1)	28 (6.0)
G2	1620 (35.0)	498 (30.7)	228 (14.1)
G3	1460 (31.6)	598 (41.0)	346 (23.7)
Unknown	1078 (23.3)	593 (55.0)	213 (19.8)
ER (oestrogen receptor) status			
Positive	2766 (59.8)	829 (30.0)	395 (14.3)
Negative	749 (16.2)	361 (48.2)	227 (30.3)
Unknown	1112 (24.0)	612 (55.0)	193 (17.4)
Surgery			
No	673 (14.6)	544 (80.8)	280 (41.6)
Yes	3331 (72.0)	918 (27.6)	454 (13.6)
Unknown	623 (13.5)	340 (54.6)	81 (13.0)
Radiotherapy			
No	1802 (39.0)	832 (46.2)	356 (19.8)
Yes	2208 (47.7)	636 (28.8)	382 (17.3)
Unknown	617 (13.3)	334 (54.1)	77 (12.5)
Chemotherapy			
No	2938 (63.5)	1097 (37.3)	463 (15.8)
Yes	1060 (22.9)	367 (34.6)	272 (25.7)
Unknown	629 (13.6)	338 (53.7)	80 (12.7)
Adjuvant endocrine therapy			
None	1364 (29.5)	576 (42.2)	302 (22.1)
Tamoxifen only	2395 (51.8)	953 (39.8)	347 (14.5)
Tamoxifen and Als	444 (9.6)	154 (34.7)	106 (23.9)
Als only	424 (9.2)	119 (28.1)	60 (14.2)

Table 1. (Continued)

	Number of women (%)	All-cause mortality (%)	Breast cancer mortality (%)
Aspirin use			
Never	3383 (73.1)	1225 (36.2)	563 (16.6)
Prediagnosis only	209 (4.5)	124 (59.3)	68 (32.5)
Pre- and postdiagnosis	473 (10.2)	249 (52.6)	108 (22.8)
Postdiagnosis only	562 (12.2)	204 (36.3)	76 (13.5)
Abbreviation: Als = aromatase inhibitors.			

with an increased risk of all-cause mortality. Increasing social class, having surgery, positive ER status and using adjuvant endocrine therapy were associated with a reduced risk of all-cause mortality (see Supplementary Table 1). However, aspirin use prediagnosis was associated with increased risk of all-cause mortality (HR = 1.62, 95% CI = 1.42–1.85, $P < 0.001$).

Cox's regression models were also used to report HRs for patients who died from breast cancer. After adjustment, the use of aspirin postdiagnosis was associated with a lower risk of breast cancer mortality (HR = 0.42, 95% CI = 0.31–0.55, $P < 0.001$). Increasing tumour stage, lymph node involvement and tumour grade, the presence of metastases as well as having chemotherapy were associated with an increased breast cancer mortality. Having surgery or radiotherapy, positive ER status and using adjuvant endocrine therapy reduced the risk of breast cancer mortality. Patients aged 50–59 years were at lower risk than those aged under 40 years, but there was no other effect by age and SES had no effect on risk of breast cancer mortality (see Supplementary Table 2). Again, prediagnosis aspirin use was associated with an increased risk of breast cancer mortality (HR = 2.10, 95% CI = 1.73–2.55, $P < 0.001$).

To examine the effect of adherence to aspirin on all-cause mortality, patients were classed as having high adherence: 80% or above (789 patients, 76% of patients taking aspirin postdiagnosis) or low adherence (246 patients, 24%) and compared with those patients not taking aspirin. Patients with low adherence had a reduced HR for death (HR = 0.56, 95% CI = 0.45–0.69, $P < 0.001$) as did for those with high adherence (HR = 0.55, 95% CI = 0.48–0.62, $P < 0.001$) compared with non-users.

DISCUSSION

Women who used aspirin following a diagnosis of breast cancer had a reduced risk of all-cause and breast cancer-specific mortality. Advancing age was related to increased risk of death attributable to increased comorbidities, such as cardiovascular disease and stroke. The younger patient age group also fared worse, in keeping with recognised outcome data (Copson *et al*, 2013).

Patients in the most affluent socioeconomic quintile had a lower risk of all-cause mortality as reported elsewhere in the literature (Fein, 1995; Mackenbach *et al*, 2003; Marmot, 2003). We found a significantly higher all-cause mortality risk in those patients who took aspirin prediagnosis, suggesting aspirin use as a marker of pre-existing cardiovascular morbidity, which may also mean that they are unable to receive optimal therapy for breast cancer.

The present population-based cohort study confirms the reduction in all-cause mortality with aspirin use (adjusted HR = 0.53, 95% CI = 0.36–0.79) (Holmes *et al*, 2010), and breast cancer-specific mortality (adjusted HR = 0.53, 95% CI = 0.30–0.93) (Holmes *et al*, 2010), or when aspirin was taken postdiagnosis 2–5 days a week (multivariate RR = 0.40, 95% CI = 0.24–0.65, $P = 0.03$) (Blair *et al*, 2007) and 6–7 days a week (multivariate RR = 0.57, 95% CI = 0.39–0.82, $P = 0.03$) (Blair *et al*, 2007).

Table 2. Characteristics of patients classified by aspirin use postdiagnosis

	Post-diagnosis non-aspirin users (%)	Post-diagnosis aspirin users (%)	χ^2 Test, d.f., P-value
Number of women	3592	1035	
Age at diagnosis (years)			
<40	205 (5.7)	2 (0.2)	400.7, 5, <0.001
40–49	608 (16.9)	33 (3.2)	
50–59	934 (26.0)	166 (16.0)	
60–69	821 (22.9)	297 (28.7)	
70–79	587 (16.4)	296 (28.6)	
80+	434 (12.1)	241 (23.3)	
Socioeconomic status (SCSIMD5)			
1 (most deprived)	450 (12.5)	149 (14.4)	50.1, 5, <0.001
2	474 (13.2)	125 (12.1)	
3	527 (14.7)	204 (19.7)	
4	1124 (31.3)	321 (31.0)	
5 (most affluent)	759 (21.1)	210 (20.3)	
Unknown	258 (7.2)	26 (2.5)	
Tumour stage			
1	808 (22.5)	224 (21.6)	13.4, 4, 0.009
2	909 (25.3)	302 (29.2)	
3	209 (5.8)	50 (4.8)	
4	235 (6.5)	87 (8.4)	
Unknown	1431 (39.8)	372 (35.9)	
Node status			
N0	2001 (55.7)	594 (57.4)	1.9, 3, 0.590
N1	392 (10.9)	119 (11.5)	
N2	79 (2.2)	21 (2.0)	
Unknown	1120 (31.2)	301 (29.1)	
Metastases			
No	2270 (63.2)	689 (66.6)	9.8, 2, 0.007
Yes	134 (3.7)	20 (1.9)	
Unknown	1188 (33.1)	326 (31.5)	
Tumour grade			
G1	347 (9.7)	122 (11.8)	9.5, 3, 0.023
G2	1235 (34.4)	385 (37.2)	
G3	1164 (32.4)	296 (28.6)	
Unknown	846 (23.6)	232 (22.4)	
ER (oestrogen receptor) status			
Positive	2139 (59.6)	627 (60.6)	1.7, 2, 0.428
Negative	595 (16.6)	154 (14.9)	
Unknown	858 (23.9)	254 (24.5)	
Surgery			
No	484 (13.5)	189 (18.3)	26.54, 2, <0.001
Yes	2651 (73.8)	680 (65.7)	
Unknown	457 (12.7)	166 (16.0)	
Radiotherapy			
No	1351 (37.6)	451 (43.6)	27.0, 2, <0.001
Yes	1787 (49.8)	421 (40.7)	
Unknown	454 (12.6)	163 (15.8)	
Chemotherapy			
No	2174 (60.5)	764 (73.8)	121.2, 2, <0.001
Yes	954 (26.6)	106 (10.2)	
Unknown	464 (12.9)	165 (15.9)	
Adjuvant endocrine therapy			
None	1165 (32.4)	199 (19.2)	74.3, 3, <0.001
Tamoxifen only	1806 (50.3)	589 (56.9)	
Tamoxifen and Als	325 (9.1)	119 (11.5)	
Als only	296 (8.2)	128 (12.4)	
All-cause mortality	1349 (37.6)	453 (43.8)	13.0, 1, <0.001
Breast cancer mortality	631 (17.6)	184 (17.8)	0.02, 1, 0.875

Abbreviation: Als= aromatase inhibitors.

Table 3. Association of aspirin use with all cause and breast cancer mortality

	Unadjusted HR (95% CI), P-value	Adjusted HR (95% CI), P-value
All cause mortality		
Pre-diagnosis aspirin use		
Never	1.0	1.0
Ever	2.08 (1.86–2.34), <0.001	1.62 (1.42–1.85), <0.001
Aspirin use		
No	1.0	1.0
Yes	1.08 (0.92–1.26), 0.358	0.53 (0.45–0.63), <0.001
Breast cancer mortality		
Pre-diagnosis aspirin use		
Never	1.0	1.0
Ever	2.08 (1.86–2.34), <0.001	1.62 (1.42–1.85), <0.001
Aspirin use		
No	1.0	1.0
Yes	1.08 (0.92–1.26), 0.358	0.53 (0.45–0.63), <0.001

The Adjusted Models allowed for age, socio-economic status, cancer stage and grade at diagnosis, ER status, surgery, radiotherapy, chemotherapy and adjuvant endocrine therapy. The complete tables are available as supplementary information online.

The study looked at all breast cancers diagnosed from an unselected population-based cohort, which included people from all levels of SES and age. The study used dispensed prescribing records from a closed prescribing system where every prescription presented to a pharmacist is subsequently recorded and collated. The prescribing data were then linked to clinical cancer records, cancer registry records and health board population databases to provide the study data set.

Limitations include the lack of over-the-counter medication information for Scotland, although previous work has suggested that long-term use of aspirin is mainly through prescriptions (Morant *et al*, 2004). We did not have recurrence information on this population as it is not robustly recorded in the routine data sets this work was based upon. Unfortunately, the indication for the aspirin prescription and the actual use (rather than filling the prescription) were not recorded. Most (99%) of the aspirin used was a low 75 mg dose, so it is unlikely to have been used as analgesia for family members. Similarly, there was no patient data available on smoking status or BMI, useful in establishing a link to increased risk of mortality when aspirin was taken prediagnosis. Long-term low-dose aspirin is associated with an increased risk of peptic ulcer and gastrointestinal bleeding, but we had no information on toxicity and side effects from aspirin use (Yeomans, 2011). It should also be noted that this was an observational study and so there is the potential for the estimated associations to be the result of unmeasured residual confounders.

Clinical implications. Our findings demonstrate a substantial risk reduction in all-cause mortality and breast cancer-specific mortality when aspirin is taken following a diagnosis of breast cancer. Because aspirin use was associated with decreased risk of death from breast cancer, this suggests that aspirin has a direct interaction with the disease, with PTGS-2 inhibition a potential molecular mechanism for aspirin to halt the growth of a tumour and prevent metastasis. It would be very interesting if the survival benefit gained is due to something as nonspecific as PTGS-2 inhibition, thus supporting the current prospective trials of aspirin in breast and other cancer types (Phillips *et al*, 2013).

Conclusions. This population-based cohort study suggests that low-dose aspirin prescribed following a diagnosis of breast cancer is associated with a decreased risk of all-cause and breast cancer-specific mortality. Further studies are needed to investigate the exact mechanism of this protective effect. Prospective randomised clinical trials may well define the effectiveness of aspirin for specific patient subgroups in the near future.

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CONFLICT OF INTEREST

All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coi_disclosure.pdf; none of the authors have financial relationships with any organisations that might have an interest in the submitted work in the previous 3 years and have no other relationships or activities that could appear to have influenced the submitted work.

ETHICS STATEMENT

All data were fully anonymised and data use were compliant with the HIC research governance process, which are approved by the Tayside Committee on Medical Research Ethics and the Caldicott Guardian.

AUTHOR CONTRIBUTIONS

CMcC conceived the study and planned it with FMS and AMT. DF carried out the analysis with support from CMcC. All authors contributed to the writing of the paper. CMcC is the guarantor.

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