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The Economics of Higher Education in Virginia

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THE ECONOMICS OF HIGHER EDUCATION IN VIRGINIA

Phillip A. Jones Clarence R. Jung 1989-1

THE ECONOMICS OF HIGHER EDUCATION IN VIRGINIA

by .

Phil Jones and Clarence Jung - 1989

Background

This study is based primarily on financial data from colleges and universities in Virginia for the two academic years 1987-88 and 1988-89. The data base stems from detailed data submitted by virtually all colleges and universities to the federal and state governments and reported in the Integrated Postsecondary Education Data survey (IPEDS). The data were accessed by Professor Jones from the State Council of Higher Education using the Bitnet computer program. (These data can be accessed for any college or university in the country, but this study employs only Virginia data.)

The aim of the study is to gain an understanding of the matrix of financial, economic and academic factors which determine the nature of a college or university; and to interpret these factors as they relate to decisions facing faculty, administrators, government officials, students and families and other interested parties. Attention to these matters in the media today would suggest that this subject is hardly an irrelevant one. General Approach

This research proceeds, broadly speaking, in two ways. First, regression analysis is used to develop single equation models of the economic and academic factors involved in the source and application of funds for higher education. The equations use SAT scores as representative of the "quality" of schools as the primary dependent variable and various independent variables such as faculty salaries, scholarship assistance, tuition, academic support among others.

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Second, ratio analysis and comparisons of revenue and expenditure patterns are employed to find methods of evaluating colleges and universities for efficiency in the use of resources. In this approach, the intent is to aply the model of corporation finance to colleges and universities. (At the time of this writing, a good bit of the work in this approach is yet to be done.)

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Tatle 1

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s = 60.51 R-sq = 89.1% R-sq(adj) = 82.0%

C1 = SAT C2 = FTE (School Size) C3 = Faculty Salary C4 = Tuition C5 = (Endowment Income)/Student C6 = Scholarsh:;s/Student C7 = (Academic Expenditures)/Student C8 = Black/White C9 = (Academic Support)/Student C10 = (Government Appropriations)/Student

Durbin-Watson Statistic = 2.62 (not significant at 5% level) Date in: ECONED.REG;6 Regression Models

Single equation regression models were developed for the private and public sectors, taken separately. Ultimately, simultaneous equation models are more appropriate (cf. Dolan, Jung and Schmidt, 1985). The data set for Virginia schools does not, however, readily lend itself to such models.

An immediate problem posed by single equation models is that this interdependence manifests itself in the condition of multicollinearity - i.e., variables can be expressed as linear functions of each other. The presence of multicollinearity is shown in a function developed for the private sector in which SAT scores are a linear function of the following factors: size of school, faculty salaries, tuition, endowment income, scholarships, academic expenditures, black/white student composition, student services expenditures, and government aid. (These variables expressed on a per-student basis where appropriate.)

The regression is shown in TAble 1, where it is clear that the high multiple R-squared and the lack of significance in the regression coefficients is the classic case of multicollinearity. This problem is addressed by selecting a few independent variables based on previous research and on observation of what appear to be important relationships among the variables of the world of Academe.

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On that basis, the following function suggests that high SATs are associated with: well paid faculty; high tuition, historically white student cody composition; and the provision of student services.

SAT = 387 + 11.2(Faculty Salary) + .04(Tuition) -82.0(Race) (7.63) (7.44) (3.16) +56.9(Student Services) - .02(Scholarships) (9.48) (1.51) Adjusted R-square = 97%

where faculty salaries are measured in thousands of dollars, tuition in dollars per student, race as zero or one for white vs black composition of student body, student services and cholarships in dollars, t-values are given in parentheses.

This equation yields results that are not counter-intuitive. Guality is associated with higher "price" (tuition), more expensive factors of production (higher faculty salaries), and the provision of amenities (student services). Scholarships do not appear as a statistically significant factor, possibly reflecting the mixture of financial resources available to students for financial aid.

However, the line item of scholarships is related to tuition, as expressed in the following function:

Scholarships = \$554 + .18(Tuition) (2.37)

Adjusted R-square = 17%

indicating that, on average, schools increase scholarship money at a rate of \$18 for every \$100 increase in tuition. (Looking at increases in tuition versus increases in scholarships for the two years, 1986-87 to 1987-88,

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It appears that the average increase is about \$18, but the measure is statistically insignificant and the correlation is near zero. Appendix sets out the tuition charges for private colleges and universities versus increases in tuition for these two years - a pattern showing little or no consistency, suggesting that these schools are hardly "colluding" on tuition increases, as has been recently alleged by government.)

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That tuition is a major explanatory variable is indicated in the following equation:

SAT = 544 + .06(Tuition) (5.17)

Adjusted R-sq = 53%

indicating that for every \$1000 in tuition, SATs rise by 60 points. Of course, as indicated above, for every \$100 increase in tuition, scholarships rise \$20. And, by the following equation

> Tuition = \$5120 + 1.27(Endowment Income) (5.68)

for every \$100 increase in endowment income, tuition charges rise by \$127. (n.b., tuition rises with endowment income rather than being "offset" by endowment income.)

It is interesting to note the role of government aid for the private sector. SATs as a function of government assistance is

SAT = 1044 - .14(\$government aid).

Government aid is defined here as assistance from all levels of government. The direction of "causation" is quite plausibly from low SATs to government aid, suggesting that for every drop of 100 points in average SAT scores, governments invest some \$700 in aid. This suggests that government aid to private schools is perhaps doing no more than barely keeping some struggling schools in existence. This, in itself, may have important social significance. φ

The composite picture which emerges from these regressions is one which shows high quality schools (as measured by average SAT scores of incomng freshmen) with high tuition, high faculty salaries, high levels of spending per student and with the level of government assistance inversely related to SATs. Other research studies indicate that alumni achievement in business and the professions is directly related to quality of student, faculty, and per capital spending (Dolan, Jung, and Schmidt, 1985). Still other research shows that SAT levels are closely correlated with family income (cf <u>The New York Review of Books</u>, October 12, 1989, p 67).

The foregoing analysis is based on standard regression analysis techniques. Another way of comparing schools is to rank order schools along dimensions suich as operating margin, increase in total revenue, return on endowment and so on. Data of this nature is given in Appendix II. Comparing the institutions which are at the polar edges of wealth, it turns out that St. Paul's and Richmond are (not surprisingly) inversely related:

which suggests that the ranking of St. Paul's can be predicted for

any item by taking 26.2 as base and subtracting .634 of Richmond's rating. Such a comparison does not yield a recipe for financial and operating management per se, but does suggest that the economic health of an institution is not achieved without large, one might say massive, does of money. It is interesting that St. Pauls and Sweet Briar are not correlated. Perhaps only the polar cases are significant, resource allocation within schools varying so much depending upon the aims, history and mission of particular schools.

(Analysis of this nature - i.e., examination of financial ratios, etc. is planned for joint work by Professors Jones and Jung. The example shown here is indicative of the nature of this work and possible conclusions which might emerge from such work.)

Analysis of public colleg's and universities yields conclusions remarkably similar to those found for the private sector. Indeed, the fiscal factors shaping higher education in the public sector appear to be broadly ther same as those for the private sector, with what seems to be one important difference. The difference is that there are political pressures for "spreading the wealth" in the public sector.

This hypothesis receives support from the equation below

which makes SATs a function of: size of school, faculty salaries, tuition, endowment income, scholarship assistance, academic expenditures, race student services and government aid.

> SATs = 661 + .004(school size) + 8.8(Faculty Salaries) (0.74) (1.10) +.03(Tuition) +.67(Endowment Income) - .56(Scholarships) (0.64) (2.29) (-2.37) -.01(Academic Expenditures) +84(Black/White) +.18(Student Services) (-0.39) (0.51) (0.39) +0.1660vernment Aid)

In a sense, these results are surprising. The only variable showing positive association with SATs is endowment income (not a surprise in itself) while the scholarship factors is negative. It is surprisng, for example, however that government appropriations are not positively related to SATs. Certainly the image is that stages spend more on the pretigious flag-ship (high SAT) schools. (Phil has some data which indicates in that they do.) The plausible explanation is that governments try to "even things out."

The two-variable regressions (i.e., regressions relating SATs, in turn, one-on-one to faculty salary, tuition,...) do indicate that, broadly speaking the factors at work in the poublic sector are the same as for the private sector. For example, SATs and faculty salaries are positively related, and statistically significant as are the SATs and tuition. However, other variables are not significant: government assistance (as in the multiple regression); student services; academic expenditures. (Appendix III) Selecting a few variables (addressing multicollinearity and inserting "judgment" into the analysis), SATs in publicly supported colleges and universities are explained well by only three variables: faculty salaries, tuition, and race. (Table 2.).

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Thus, it appears that state policy is to invest in higher faculty salares in the flagship schools and that these schools charge higher tuition, pay faculty well, and over the years have accumulated endowment resources. These factors reenforce the already-established high quality of the institutions. However, when it comes to scholarships, academic support, and student services - the policy of the state government seems much more egalitarian. Indeed, the two-variable regression relating SAT and scholarships indicates that scholarship monies are "spread" across institutions in such a way that there is no differential impact among schools.

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The regression equation is C1 = 391 + 13.8 C3 + 0.0505 C4 - 331 C8 Predictor Cour Stdev L-Tetio Constant 391.3 152.4 2.57 0.026 $\mathbb{C}\mathbb{R}^{n}$ 2.76 04 0.05047 0.03074 1.64 0.129 58 -031.42 -6.27 0.000 s = 65.29R−sq = 89.6% R−sq(adj) = 86.8% Assign to the English SOURCE DF SS MS F Regression 3 403820 134607 31.57 0.000 Error 11 46894 4263 Total 14 450714 SOURCE DF SEQ SS CЭ 4

> C1 = SATC3 = Faculty Salary C4 = TuitionC8 = Black/White Student Body

> > All variables statistically significant at 5% level. Durbin-Watson = 2.57 (n = 24)

Table 2

A Broader Perspective

The picture which emerges from this analysis is clear. It shows that the quality of an institution depends on the dollars available to it and that these dollars, in turn, produce quality.

Hardly a startling discovery. However, the specificity of the analysis is interesting. While the quality of a school is a function of many factors, it turns out that just a few variables are the crucial ones. These are: tuition, faculty salaries, and endowment income.

Interestingly, scholarship expenditures are either statistically insignificant or are actually inversely related to quality. A plausible explanation for this finding is that the social contract in today's world is one in which the aim - both in the private and public sectors - is to make family income neutral in terms of access to higher education.

While the analysis in this paper, and the analysis in the literature on this subject, find a number of other relationships that are of interest (for example, spending for academic support is statistically significant in the private, but not the public sector), the nub of the matter is that price, income and quality me inextricably entertwined.

Of course, this is the way the world works and that this should be true in higher education is not (as indicated above) surprising. However, the implications are not trivial when viewed in a broader context. That context is that graduates of quality schools go on to high levels of achievement in the business and professional world and thus are in a position to ensure alma mater's continued success and alma mater's continued ingestion of students whose socio-economic background prepares them well for entrance to prestigious schools and whose family income levels permit the payment of the high cost of higher education.

Thus, the findings of this study might be summed up in the phrase from the popular song of the 1920s (Ain't We Got Fun) that "the rich get rich and the poor get poorer." One might interpret this as evidence of the efficient working of a market system in which innate personal ability and a supportive family background lead to high levels of output and productivity. Or, one might interpret this as evidence for the Marxist view that the whole education system of a society is merely an instrument for the fashioning of a subservient work force.

That the graduates of quality schools do better in the business and professional world is shown in a study by Dolan, Schmidt and Jung (1985, Review of Economics and Statistics) in which a simultaneous equation model was developed showing the interdependence of student ability, faculty salaries and alumni achievement and the role of various exogenous factors. The focus of that study was on the identification of patterns of resource allocation within a school that would produce successful alumni. The study concluded that "faculty salary, academic and administrative suppoprt ...quality students and quality faculty, buttressed by ...libraries, laboratories, and, more recently, computers, appear as the major cogs driving the educational process" pp. 519-520

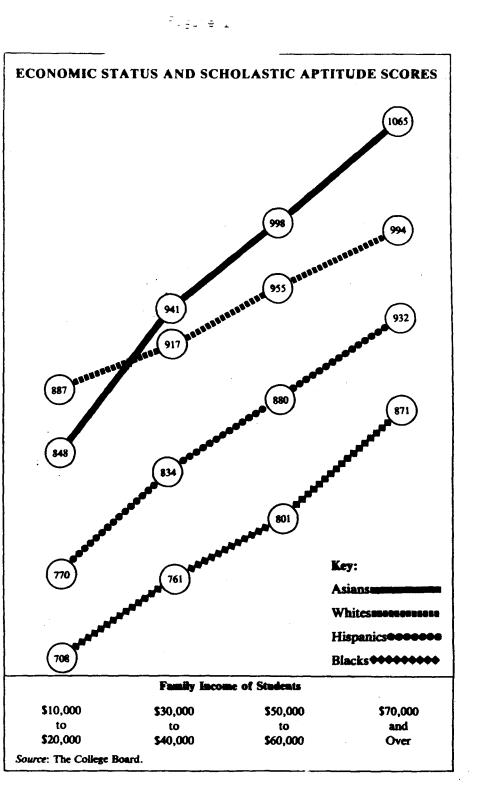
Although the data base for Virginia schools is more limited in this respect than for the Dolan-Jung-Schmidt study, analysis of this data set via two-variable regressions indicates that the production of Ph.D.s and Executives are related as follows:

> Executives = -6.99 + .00843(SAT) (3.92) R-square = 39% Ph.D.s = 26.2 + .033(SAT) (3.52) R-square = 33%

That the quality of students appearing on these campuses is a function of family income and socio-economic status has been documented by the College Board, and set out at some length in a recent review article in the <u>New York Review of Books</u>. This is graphically illustrated in Figure <u>I</u>). Taking the values in this graph and performing two-variable regression analysis indicates that SATs are a statistically significant function of income. The graph shows unambiguously, also, that scores on the SAT tests are a function of ethnic background.

That financially disadvantaged students lack access to higher education is not clear. The results in this paper indicate that there is an inverse relationship between quality of students and scholarship aid (or that the relationship is not significant). The Dolan <u>et al</u> study found the coefficient for the scholarship variable was negative (and significant at the .01 level). However, at least one study (Machlis, circa 1974) found that

low-income classes are underrepresented in higher education and that



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the wealthy have a "disproportionately large number of students in attendance."

Policy Implications

The crystal-clear indications are, from this study, that higher education in Virginia is a product of, and a component of, the social and economic system. That this statement is tinged with an economic interpretation of history (Marxist, to some degree) should not blind one to its legitmacy and importance.

It suggests that, in the first place, that the system has

worked well in the past and serves society well in the present. To use a perhaps tired cliche, whatever is not broken should not be fixed. However, to say that the system works generally well and effectively is not to say that there are not important possibilities for change.

One important change is suggested and that is that consideration should be given to much higher levels of spending for low income and minority groups. The results of the present paper indicate that there is a definite thrust toward financial assistance to these groups. The algebraic signs of the coefficient for scholarship money and for government aid to private institutions would suggest this. Also, the apparent "spreading" of financial resources among the public institutions warrants such an inference.

But the overriding evidence here is that this is only marginally effective. The implication is that the term "massive" might be the operative term. Large doses of capital from the private and public sectors might be in order. Much of that capital would be most effectively used at the elementary and secondary school levels (see, e.g., the Dolan and Schmidt study, 1987, <u>Economics of Education</u> <u>Review</u>). Also, one might argue that the pricing system in higher education should, in theory at least, involve even more price discrimination than is presently the case.

While these conclusions would appear to have a considerable support in the context of this paper, an even broader context would suggest that education is not the only scarce resource in society and that spending for health, transportation, corrections, defense, recreation might create an opportunity cost that would preclude

higher spending for education. The general equilibrium analysis required to address this matter is beyond the scope of this study.

At the level of partial equilibrium analysis, however, it is clear that quality, price and income are the key determinants of the nature of higher education.

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APPENDIX I

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Curbin -Watson statistic = 1.61

APPENDIX II

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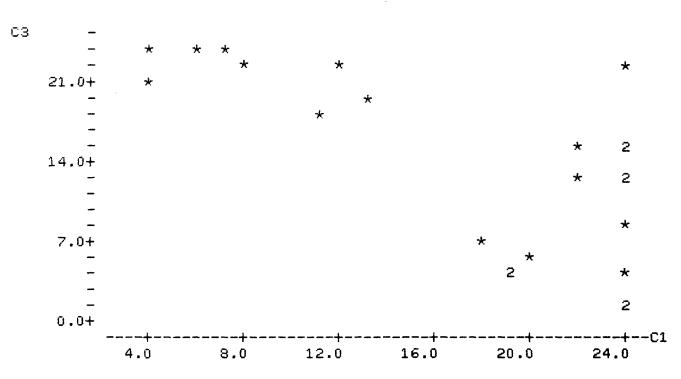
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St Paul's is UgR

3C1 = 26.2 - 0.634 C3

Predict Constan C3	it 🔨	Coef 26.169 3.6337	Stde 2.37 0.150	5 11	tio .02 .23 √	
s = 5.6	507	R-sa = 45	. 9%	R−sq(adj)	= 43.4%	
Analysi	s of Van	riance				
SOURCE	DI	-	SS	MS		
Regress	ion 1	561	.33	561.33		
Error	2:	L 660	.32	31.44		
Total	22	2 1221	.65			
Obs.	C3	C1	Fit		Residual	St.Resid
1	15.0	22.00	16.66	1.18	5.34	0.97
1 2 3	7.0	18.00	21.73	1.55	-3.73	-0.69
3	9.0	24.00	20.47	1.37	3.53	0.65
4	12.0	24.00	18.56	1.20	5.44	0.99
5	4.0	24.00	23.63	1.88	0.37	0.07
6	5.0	20.00	23.00	1.76	-3.00	-0.56
Continu MTB > p						

St Paul's is Ug R





Va Union on U of R

4 24 15 12 1 MTB > regress c4 on 1 predictor in c3 The regression equation is C4 = 18.8 - 0.307 C3Predictor Coef Stdev t-ratio ~18.838 3.153 5.97 Constant \checkmark СЗ -Ø.3069 0.1991 -1.54s = 7.445R-sq = 10.2% R-sq(adj) = 5.9% Analysis of Variance SS 131.63 SOURCE DF MS 131.63 Regression 1 Error 21 1163.85 55.42 Total 22 1295.48

MTB >

Va Union on Suret Brian

MTB > regres	s c4 on	1 predictor	in c2	
The regressi C4 = 9.17 +				
Predictor Constant C2 (Coef 9.173 +0.3149	4.	dev 703 566	t-ratio 1.95 1.23
s = 7.587	R-sq	= 6.7%	R-sq(a	dj) = 2.2%
Analysis of	Variance			
SOURCE Regression Error Total	DF 1 21 22	SS 86.70 1208.77 1295.48	86. 57.	

MTB >

Va Union on St Paul's.

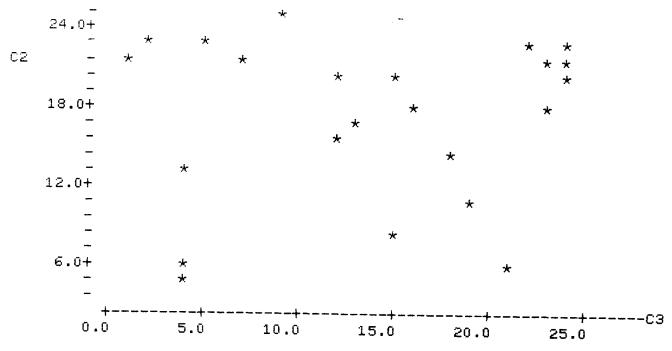
The regression equation is C4 = 7.99 + 0.380 C1	
Predictor Coef Stdev t-ratio Constant 7.988 3.947 2.02 C1 (+) 0.3797 0.2089 1.82	\checkmark
s = 7.301 R-sq = 13.6% R-sq(adj) = 9.	5%
Analysis of Variance	
SOURCEDFSSMSRegression1176.17176.17Error211119.3153.30Total221295.48	
Unusual Observations Obs. C1 C4 Fit Stdev.Fit Res 4 24.0 1.00 17.10 2.05 - R denotes an obs. with a large st. resid.	

MTB >

Server Brian is U of R

C2 = 16.0 + 0.095 C3								
Predictor Constant C3	Coef 15.956 0.0946		tio .88 .55					
s = 6.407	R-sq = 1.4%	R-sq(adj)	= 0.0%					
Analysis of Variance								
SOURCE Regression Error Total	DF 53 1 12.52 21 861.93 22 874.43	12.52 41.04						
Obs. C3 1 15.0 2 7.0 3 9.0 4 12.0 5 4.0 6 5.0 Continue? n MTB >	22.00 1 25.00 1 15.00 1 13.00 1	Fit Stdev.Fit 7.38 1.35 6.62 1.77 6.81 1.57 7.09 1.37 6.34 2.14 6.43 2.01	Residual -9.38 5.38 8.19 -2.09 -3.34 6.57	St.Resid -1.50 0.87 1.32 -0.33 -0.55 1.08				

Sucert Brien vs UdR.

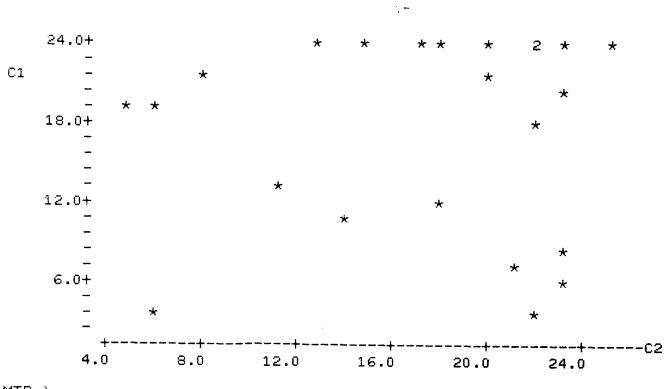




St Paul's Vs Savert Brian.

The regression equation is C1 = 16.4 + 0.061 C2 Bonglightr 16.381 4.721 t-ration C2 0.0611 0.2576 0.24 /								
s = 7.617 R-sq = 0.3% R-sq(adj) = 0.0%								
Analysis of Variance								
SOURCE Regression Error Total		SS 3.26 1218.39 1221.65	MS 3.26 58.02					
Obs. C 1 8. 2 22. 3 25. 4 15. 5 13. 6 23. Continue? n	0 22.00 0 18.00 0 24.00 0 24.00 0 24.00	16.87 17.72 17.91 17.30 17.17	Stdev.Fit 2.87 2.00 2.55 1.69 1.93 2.17	Residual 5.13 0.28 6.09 6.70 6.83 2.21	St.Resid 0.73 0.04 0.85 0.90 0.93 0.30			

St Paul's vs Suret Brian



 $MTB \rightarrow$

8/14/89

Correlation matrix

2 3 4 7 5 6 7 1 × . 42 .38 .52 -.06 - 0.77 .72 2 ,42 × 195 .81 .10 -.05 •73 3 ,38 ,95 × .72 ,2/ -.13 .62 ¥ 181 172 × .52 .01 -.05 .67 Ś -.06 121 ,01 -107 .01 , 10 \mathbf{x} 6 -.05 -.13 -.05 -.07 -77 × -129 7 .72 .72 162 .67 .01 -.29 × 1= SAT 2= Academin # / FTES 3 = Total \$/FTES 4 = Trution #/FISS 5= Just # / FTES

i t

> 6 = Llemy white Black 7 = Avg Faculty #

04 05 07 08 09 010	$\begin{array}{c} 0.516 \\ 0.510 \\ -0.362 \\ 0.415 \\ -0.774 \\ 0.132 \\ 0.211 \end{array}$	0.345 0.216 0.119 0.533 -0.212 -0.230 0.668	0.674 0.609 0.177 0.739 -0.289 0.115 0.627	0.651 0.381 0.814 -0.054 0.386 0.735	0.521 0.741 -0.177 0.635 0.476	0.568 0.668 0.464 0.596	-0.053 0.429 0.882	-0.000 0.112
	C9							
C3.0	tri postiki							
мтв >	corr betwee			05		C7		60
~ ~	C1	СЗ	C4	C5	C6	1.7	C8	C9
					00	0.		0.5
C3	0.723	0 674			00	0.		05
C4	0.516	0.674	0 651		00	0.		05
C4 C5	$\begin{array}{c} 0.516 \\ 0.510 \end{array}$	0,609	0.651 0.381			0.		05
C4	0.516		0.651 0.381 0.814	0.521 0.741	0.568	0.		0.5
C4 C5 C6	0.516 0.510 -0.362	0.609 0.177	0.381	0.521	0.568 0.668	-0.053		0.5
C4 C5 C6 C7	0.516 0.510 -0.362 0.415	0.609 0.177 0.739	0.381 0.814	0.521 0.741	0.568		-0.000 0.112	0.243

 $\mathsf{MTB} \rightarrow$

APPENDIX III

Appendix III

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Size

C1 = 882 + (C1 = 882 + 0.0113 C2							
Predictor Constant C2		Stdev 71.89).006568	t-ratio 12.26 1.71					
s = 168.2 R-sq = 18.4% R-sq(adj) = 12.2%								
Analysis of	Analysis of Variance							
Sfreesion Regression Error Total	DF 1 83(13 367(14 450)	539	MS 83075 28280	F 2.94	р 0.110			
055. 01 1 671: 2 323. 3 293: 4 975: 5 1282: 6 70: 0::tinje? n 75:	6 1240.0 2 939.0 6 915.0 4 1095.0 0 950.0	957.3 918.0 914.7 991.5	45.4 56.5 57.7 43.9	103.5 -76.0	1.75 0.13 0.00 0.64 -0.47			

The regression equation is C1 =10 + 25.7 C3

Et i est	Tana a S	Stdev	t-ratio	Ę
Constant	9.9	259.3	0.04	0.970
Ċ.B	25.738	6.823	3.77	0.002

s = 129.7 R-sc = 52.3% R+sq(adj) = 48.6%

Analysis of Variance

•

SCURCE Regres Error Total		DF 1 13 14	21	SS 5516 5197 0714	MS 235516 16554 (~	F 14.23	p 0.002
Obs. 1 2 3 4 5 6	C3 44.4 33.7 33.0 35.6 38.5 35.2	1	C1 240.0 939.0 915.0 095.0 950.0 914.0	Fit 1152.7 877.3 859.2 926.2 1000.8 915.9	Stdev.Fit 56.6 42.9 46.1 36.1 33.7 37.3	Residual 87.3 61.7 55.8 168.8 -50.8 -1.9	St.Resid 0.76 0.51 0.46 1.37 -0.41 -0.02
Contin	ue:						

Formet 9

Turtion

The regression equation is C1 = 713 + 0.118 C4							
Predic Constan C4	nt	Coef 712.7 J.1 And	7 129.8	5.49	0.000		
s = 15:	9.5	R-sq	1 = 26.6%	R-sq(adj)	= 21.0%		
Analys.	is of '	Variance	2				
SOURCE Regress Error Total		DF 1 13 14	SS 119916 330798 450714	MS 119916 25446	F 4.71	р 0.049	
Obs. 1 2 3 4 5 6 Continu	C4 3143 1662 1654 1887 2079 1350	1240 939 915 1090 950 914	10 1083.8 10 908.9 10 908.0 10 935.5 10 958.2	52.6 52.8 46.0 42.4	156.2 30.1 7.0 159.5	1.07 0.20 0.05 1.04	

Continue?

Scholansliges

C1 = 1078 - 0.190 C6

Predictor Constant C6	10	Coef 78.16 .1896	Stdev 83.23 8.1353	12.95	0.000	
s = 173.5		R-sq	= 13.1%	R-sq(adj)	= 6.4%	
Analysis	of Var	iance				
SOURCE	DF		SS	MS	F	Р
Regressio			59164	59164	1.96	0.184
Error	13		391550	30119		
Total	14		450714			
Obs.	C6	C	:1 Fit	Stdev.Fit	Residual	St.Resid
1	259	1240.	0 1029.1	56.9	210.9	1.29
2 3	281	939.	0 1024.9	55.1	-85.9	-0.52
3	368	915.	0 1008.4	49.2	-93.4	-0.56
	193	1095.	0 1041.6	62.8	53.4	0.33
	387 -	950.	0 1004.8	48.2	~54.8	-0.33
6	262	914.	0 1028.5	56.7	-114.5	-0.70
Continue? MTB >	n					

Endernt.

	The regression equation is C1 = 937 + 0.458 C5							
Predicto Constant C5		Coef 936.62 0.4581	46.02	20.35	p 0.000 0.052			
s = 160.1	1	R-sq	= 26.0%	R-sq(adj)	= 20.3%			
Analysis	of V	ariance						
SOURCE Regressi Error Total	on	13	58 117351 333363 450714	MS 117351 25643	F 4.58	р 0.052		
0bs. 1 2 3 4 5 6	10 0		.0 990.2 .0 941.2 .0 936.6 .0 948.1 .0 941.2	46.0 43.9 45.1	249.8 -2:2 -21.6 146.9	1.62 -0.01 -0.14 0.95		

Continue?

academie

C1 = 785 + 0.0359 C7

Philippier Constant C7	Coef 784.7 9.03587	126.6	6.20	р 0.000 0.124	
s = 169.4	R-sq	= 17.2%	R-sq(adj) =	10.8%	
Analysis of	Variance				
SOURCE Regression Error Total	DF 1 13 14	SS 77488 373226 450714	MS 77488 28710	F 2.70	р 0.124
05s. c 1 627		C1 Fit .0 1009.9	Stdev.Fit 47.4	Residual 230.1	St.Resid 1.41

1	6277	1240.0	1009.9	47.4	230.1	1.41
2	3262	939.0	901.7	64.6	37.3	0.24
3	4063	915.0	930.5	53.1	-15.5	-0.10
4	3901	1095.0	924.7	55.2	170.3	1.06
5	4641	950.0	951.2	47.1	-1.2	-0.01
6	3559	914.0	912.4	60.0	1.6	0.01
Contir	ue?n					
MTB >						

Predictor Constant C8	Coef 1032.46 -394.46	Stdev t-rati 32.73 31.5 89.63 -4.4	5 0 .000	
s = 118.0	R-sq = 59.	8% R-sq(adj)	= 56.8%	
Analysis of	Variance			
SOURCE Regression Error Total	DF S 1 26970 13 18100 14 45071	7 13924	F 19.37	р 0.001
054. 08 1 0.00 2 0.00 3 0.00 4 0.00 5 0.00 6 0.00 Continue? n MTB >) 1240.0) 939.0) 915.0) 1095.0) 950.0	Fit Stdev.Fit 1032.5 32.7 1032.5 32.7 1032.5 32.7 1032.5 32.7 1032.5 32.7 1032.5 32.7 1032.5 32.7	207.5 -93.5 -117.5 62.5 -82.5	1.83 -0.82 -1.04 0.55 -0.73

01 = 1032 - 394 08

Race

The regression equation is C1 = 895 + 0.201 CS							
Precictor Constant C9	Coef 895.4 0.2008	Stdev 182.0 0.4176	t-ratio 4.92 0.48	p 0.000 0.639			
s = 184.6	R-sq	= 1.7%	R-sq(adj)	= 0.0%			
Analysis of	Variance						
SOURCE Regression Error Total	DF 1 13 14	SS 7875 442839 450714	MS 7875 34065	F 0.23	р 0.639		
055. 05 1 441 2 446 3 386 4 284 5 385	1240 939 915 1025	.0 984.0 .0 985.0 .0 973.3 .0 <u>952.5</u>	Stdev.Fit 48.4 48.8 49.6 74.3 50.5	Residual 256.0 -46.0 -58.3 142.5 -29.4	St.Resid 1.44 -0.26 -0.33 0.84 -0.19 -0.19		

Continue?

Services

Cl = 397 + 0.0156 Cl0							
^c tedicto Constant C10	Coef 897.1 0.01558	116.0	7.73	0.000			
s = 182.0	R-sq	= 4.5%	R-sq(adj)	= 0.0%			
Apalysis of	Variance		·,				
SOURCE Regression Error Total	DF 1 13 14	SS 20155 430559 450714	MS 20155 33120	F 0.61	р 0.449		
Obs. C10 1 4872 2 3047 3 3863 4 3360 5 3872 6 3075 Continue? n MTB >	1240 939 915 1095 950	.0 973.0 .0 944.6 .0 957.3 .0 949.5 .0 957.5	55.2	267.0 -5.6 -42.3 145.5 -7.5	1.52 -0.03 -0.24 0.85 -0.04		

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