Applied Urban Fire Department Incident Forecasting

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Abstract-Every day, when firefighters respond to emergencies, they and the public face an unnecessary risk due to inadequate staffing. Having too many people stand-by costs a lot of money, on the other hand, having too few people stand-by leads to unnecessary safety risks. Therefore, for adequate staffing purposes, forecasting the number of incidents that each fire station has to handle is a very relevant question. In this paper, we develop models to create a good forecast for the number of incidents that each fire station in Amsterdam-Amstelland has to handle. Previous studies mainly focused on multiplicative models containing correction factors for the weekday and the time of the year. Our main contribution is to incorporate the influence of different weather conditions in the categories of wind, temperature, rain, and visibility. We show that an ensemble model has the best predictive performance. Rain and wind typically have a strong linear influence, while temperature mainly has a nonlinear influence.

Keywords-incident forecasting; fire department planning; generalized linear models; ensemble models.

I. INTRODUCTION

As for most organizations, the ability to accurately forecast demand is of "paramount importance" for emergency services, fire departments included [1]. In the 1970s, the Fire Department of the City of New York and The New York City-RAND Institute jointly conducted various groundbreaking studies [2]. More recent academic interest seems to be focused more on ambulance services. While there are obvious similarities between emergency service providers, they differ in (the number of) incident types, demand characteristics, and operational logistics.

Nevertheless, the problems that fire departments have to deal with, like loss of coverage and the degradation of response times, are similar. The same is true for possible gains. On a strategic and tactical level, improved forecasting of workload leads to a better placement of base stations, and improved staffing and scheduling. On an operational level, one may proactively relocate units to maximize coverage and minimize response times during major incidents [3]. All things considered, efficient planning of emergency service resources is considered crucial.

Demand is an important factor when models are being developed to improve the performance of emergency service providers. It is, however, not uncommon that, for instance, call arrival rates are estimated using ad-hoc or rudimentary methods such as averages based on historical data [4]. This may ultimately lead to a degradation of performance, or overor under-staffing [5]. In most cases, reducing response times is an important performance measure since this increases the survival rate of victims [6][7].

Numerous papers have been written on the subject of forecasting forest or wildfire occurrences, many of those using weather variables and vegetation types as part of their model [8]. Forest fire forecasting is no longer a study in academia alone. In fact, in the United States, e.g., the National Interagency Coordination Center operates a predictive service which provides decision support to the U.S. Forest Service, which facilitates pro-active management and planning of fire assets on both operational and tactical levels [9].

Although the scale of wildfire occurrences in the Netherlands is smaller than in other parts of the world, it is mainly the greater interrelationship of different types of infrastructure, i.e., the wildland-urban interface, that causes concern and even lead to surface fuel models for the Netherlands [10]. For a more urban environment, like the conurbation of Western Holland, which also includes Amsterdam, forest fire occurrences are not very common.

The occurrence of certain types of incidents which fire departments in urban settings typically respond to also correlate with weather conditions. As such, incorporating this information into the planning process of emergency services yields important advantages over current practice. Typical weather and storm-related incidents that fire departments in the Netherlands respond to are fallen trees, potentially falling debris that needs securing (roofs, construction work, scaffolding), and water damage. Another important factor is that the weather also impacts fire department operations by overwhelming available resources.

At least in the Netherlands, to the best of our knowledge, there are no known applications of forecasting algorithms that are used in practice at fire departments, being urban or specialized forest services. Given this, we aim to provide an easily applicable model that can be put to use for an urban fire department. Therefore, we quantify and model the gut feeling, which tells firefighters that on stormy days they will have busy days. The organization of this paper is as follows. In Section II, we describe the data used to obtain the forecasts. Section III describes the models used for forecasting. In Section IV, we analyze the performance of the models and state the insights. Finally, in Section V, we conclude and address a number of topics for further research.

II. DATA

The available data contains one row for each incident that happened in the region Amsterdam-Amstelland from January 2008 up until April 2016. The most interesting information includes the incident's start- and end time, location, incident type, the concerned fire station, and the number of fire trucks used. Since the size of incidents matters for the number of people you need, the focus is on forecasting the number of trucks needed.



Figure 1. Histogram of the number of fire trucks per incident.

Figure 1 shows a histogram of the number of trucks per incident. The vast majority of the incidents require only one or otherwise just a few trucks. Therefore, it makes sense to distinguish between 'big' and 'small' incidents. Big incidents are mostly due to coincidences that are hard to predict. Specifically, they do not rely on bad weather conditions or a particular time of the year in the Netherlands, for example, as with forest fires in countries with a tropical climate. This arouses the expectation that the inter-incident times of big incidents can be modeled as a Poisson process.

To test the Poisson assumption, we apply the Kolmogorov-Smirnov (KS) test on the inter-incident times in cases when more than k trucks are needed for several values of k. The KS-test shows that if we define an incident as 'big' when at least k = 6 trucks are used, then the KS-test does not reject exponentiality of the inter-incident times (approximate p-value = 0.429). However, for values of k < 6, the KS-test doubts (or rejects) this exponentiality (approximate p-value = 0.073 and 0.002 when at least k = 5 and k = 4 trucks are used, respectively. Hence, according to this result, we define an incident to be big when at least 6 trucks are needed.

Now that big incidents can be modeled by a Poisson process, it is time to focus on the small incidents. The small incidents are probably easier to predict, since bad weather conditions often cause many *small* incidents to happen (like

fallen trees, water damage, or police/ambulance assistance at traffic accidents). To study this, we first omit all incidents on December 31 and January 1. There are extremely many incidents around New Year's Eve, mainly caused by accidents involving fireworks. These conditions do not occur in the rest of the year, so it seems logical to analyze these days separately.



Figure 2. Seasonal patterns: the given percentages represent relative differences with respect to the average (in blue).

There are clear seasonal patterns in the data for the number of trucks needed throughout each year, week, and day. The plots in Figure 2 illustrate this. The pattern in Figure 2c depicts the activity cycle that an average person goes through every day of the week. The week pattern (Figure 2b) differs per type of incident and looks a little different throughout the year. The pattern in Figure 2a can be included in the model in a more subtle way than taking factors per month. The problem here is that, for instance, the differences between the beginning and end of January are considerable. We correct for this by using a Loess-smoothed function over the factors per week. We will include all these patterns in our model, which will be discussed in the next section.

Besides the time-dependent components, we want to know which weather variables we must include in our model. Therefore, we use the Pearson correlation test to determine which weather conditions have a significant influence on the number of trucks we need. The results of these tests are summarized in Table I.

TABLE I. PEARSON'S PRODUCT-MOMENT CORRELATION TESTS BETWEEN SOME WEATHER VARIABLES AND THE NUMBER OF TRUCKS USED FOR SMALL INCIDENTS PER DAY.

Category	Variable	p-value	Correlation
	Average wind speed (FG)	$< 10^{-12}$	0.132
Wind	Maximum hourly mean wind speed (FHX)	$< 10^{-15}$	0.177
	Maximum wind gust (FXX)	$< 10^{-15}$	0.189
Temperature	Average temperature (TG)	0.6897	0.007
remperature	Boolean: 1 if average > 0 (TG>0)	$< 10^{-8}$	0.105
	Rainfall duration (DR)	0.0004	0.061
Rainfall *	Total rainfall (RH)	$< 10^{-15}$	0.151
	Maximum hourly rainfall (RHX)	$< 10^{-12}$	0.132
Waihilita **	Minimum visibility (VVN)	0.2217	-0.014
visibility **	Boolean: 1 if minimum < 200m (VVN<2)	0.2893	0.010

* In 0.1 mm and -1 for <0.05 mm; ** On 0-89 scale, where 0: <100 m, 89: >70 km.

We can see from this that the minimum visibility and the average temperature both have no significant (direct) influence. However, if we consider a variable indicating whether it was on average freezing on that day, then this does have predictive value. Obviously, we also have to include some variables indicating the amount of rainfall and wind. However, the variables within these categories are highly correlated (sample correlation around 0.9) and, therefore, we may exclude some of them to simplify our model.

III. MODELS

In this section, we will create a model that predicts directly the number of trucks that each fire station needs. In the previous section, we have shown that the big incidents (with at least six trucks needed) are very hard to predict and that we can best model them by an (inhomogeneous) Poisson process. We also showed that the daily pattern of the number of trucks used for small incidents is quite standard. So, if we know for some day how many trucks are needed in total, we can quite accurately extract from this how many trucks are needed per hour. Therefore, we will try to forecast the number of trucks needed per day per fire station.

In total, we have 9 different incident clusters or types in our dataset, some of which occur much more/less often than others. In Table II, we show the correlation with respect to one variable of each four weather categories. Looking at these correlations in detail, we can see that these are often in line with our expectations. For instance, high wind speed and rainfall obviously increase the number of incidents due to 'storm and water damage' (type 9) and decrease the likelihood of 'outside fires' occurring (type 1).

We will estimate, for each incident type t, a model that predicts the number of trucks used for *small* incidents $y_{t,d}$ on date d, i.e.,

$$y_{t,d} = f_{t,d} \cdot g_{t,d} \cdot x_{t,d}$$

Here, $f_{t,d}$ is a correction factor for the week number based on a Loess-smoothed function as in Figure 3, and $g_{t,d}$ is a weekday

TABLE II. INCIDENT CLUSTERS AND CORRELATION WITH RESPECT TO WIND SPEED, TEMPERATURE, RAINFALL, AND VISIBILITY.

Cluster	Туре	Wind	Temp.	Rain	Visib.	# p/day
1	Outside fire	-0.135	0.09	-0.193	0.075	3.46
2	Animal in water Animal assistance Person in water Locked out	-0.088 -0.072 -0.041 -0.006	0.134 0.129 0.056 0.159	-0.058 -0.088 -0.023 -0.043	0.013 0.069 0.009 0.062	1.65
3	Contamination / nuisance	-	-0.228	0.038	-0.111	2.52
4	Locked in elevator Automated alarm		-0.088 -0.069	0.021 0.051	-0.015 -0.037	8.16
5	Fire rumor Inside fire General assistance water	- - -	-0.103 -0.038 -0.019	- - -	- - -	3.57
6	Police assistance	0.048	-0.062	0.026	-	1.34
7	Ambulance assistance Vehicle in water Reanimation	- - -	-0.065 -0.042 -0.086	- -	-0.039 -0.025 -0.008	8.55
8	General assistance	0.063	0.079	0.057	0.052	2.28
9	Storm- and water damages	0 3 1 9	0.028	0 279	-	2.10



Figure 3. The year pattern per week (in black) together with its Loess-smoothed variant ($\alpha = 0.3$).

factor as in Figure 2b. Both are computed separately for each incident type. Finally, the term $x_{t,d}$ contains all remaining information. This includes the average level, dependencies on the weather, a possible trend and dependencies on all other variables that we are currently not considering, but which do exist in reality.

A. Linear regression model

The first attempt to model $x_{t,d}$ is by means of the linear regression model (LM)

$$x_{t,d} = \beta_0 + \beta_1 \cdot d + \beta_2 \cdot \text{windspeed}_d + \beta_3 \cdot \text{temperature}_d + \beta_4 \cdot \text{rainfall}_d + \beta_5 \cdot \text{visibility}_d + \epsilon_{t,d},$$

where $\epsilon_{t,d}$ is assumed to have expectation zero and some finite variance. Note that this model includes an intercept (β_0), a linear trend ($\beta_1 \cdot d$) and (at most) four weather variables.

B. Generalized Linear Model

Our second model, a Generalized Linear Model (GLM) arises from an observation that the largest outlier neither has the highest wind speed nor the most rainfall. However, the *combination* of wind and rainfall might be the cause. It may, therefore, be a good idea to include also cross-effects in our

model, i.e.,

$$\begin{aligned} x_{t,d} &= \beta_0 + \beta_1 \cdot d + \beta_2 \cdot \text{windspeed}_d + \beta_3 \cdot \text{temperature}_d \\ &+ \beta_4 \cdot \text{rainfall}_d + \beta_5 \cdot \text{visibility}_d \\ &+ \beta_6 \cdot \text{windspeed}_d \cdot \text{temperature}_d \\ &+ \beta_7 \cdot \text{windspeed}_d \cdot \text{rainfall}_d \\ &+ \beta_8 \cdot \text{windspeed}_d \cdot \text{visibility}_d \\ &+ \beta_9 \cdot \text{temperature}_d \cdot \text{rainfall}_d \\ &+ \beta_{10} \cdot \text{temperature}_d \cdot \text{visibility}_d \\ &+ \beta_{11} \cdot \text{rainfall}_d \cdot \text{visibility}_d \\ &+ \epsilon_{t,d}. \end{aligned}$$

Here, $\epsilon_{t,d}$ is again a residual term with zero expectation and some finite variance. Note that this is not a GLM as one may know from the literature. The only feature that causes it to be generalized is that it now also handles the cross-term relations between the weather variables. We could have called it an *expanded* linear model as well.

C. Random Forests

The Random Forest (RF) algorithm is a machine learning algorithm that can be used for both classification and regression tasks. Compared to LM and GLM it has a large computation time, but RF is often used in practice since it generally has great performance. It will, therefore, be worth a try to implement this algorithm for our regression problem.

As input, the algorithm needs a $T \times (K + 1)$ -matrix with K explanatory variables and one observation variable (in this case $x_{t,d}$), all of sample size T. In the first iteration of the algorithm, a sample of size T is drawn with replacement from the input matrix. On this sample, a decision tree (DT) algorithm is executed. This procedure is repeated N times, yielding N decision trees. When a new sample comes in, we can take all N predictions for this sample and average these to get the final prediction.

D. Performance measures

To evaluate the different models, we create a train and a test set. The train set contains all data up until 2015/06. The test set contains all data from 2015/07 onwards. This holds for all incident types, so all test sets contain exactly nine months of data and the quality of the forecasts can, therefore, be compared easily. We will measure the quality of a forecast on n samples using the Mean Absolute Percentage Error,

$$\mathsf{MAPE} = \frac{1}{n} \sum_{t=1}^{n} \left| \frac{y_t - \widehat{y}_t}{y_t} \right| \stackrel{(y_t \ge 0)}{=} \frac{1}{n} \sum_{t=1}^{n} \frac{|y_t - \widehat{y}_t|}{y_t},$$

as well as its weighted version, i.e.,

wMAPE =
$$\frac{\sum_{t=1}^{n} \frac{|y_t - \hat{y}_t|}{y_t} y_t}{\sum_{t=1}^{n} y_t} = \frac{\sum_{t=1}^{n} |y_t - \hat{y}_t|}{\sum_{t=1}^{n} y_t}$$

Here, y_t is the true value in time period t and \hat{y}_t is the prediction.

IV. RESULTS

In this section, we will compare the performance of the different models and evaluate the insights derived from them. The results on the MAPE and wMAPE values are given in Table III. These performance measures are based on the total daily number of trucks used for small incidents (over all fire stations and types). This enables us to compare all models through one value. It is also interesting to see how significant a parameter is on a 1 to 5 scale, as in Table IV for LM, Table V for GLM, and Table VI for RF. Here, we assign 1 when the p-value < 0.001 (very significant) until 5 when the p-value ≥ 0.1 (not significant).

TABLE III. PERFORMANCE MEASURES OF THE MODELS.

Model	MAPE	WMAPE
LM	0.1886	0.1924
GLM	0.1865	0.1880
RF	0.2006	0.2019

A. Linear regression model

For the linear model, comparing Table IV to Table II, we observe that when a weather variable has significant predictive power for some type, then their mutual correlation is relatively high as well. This is a nice result, but unfortunately, the reverse is not true. For instance, type 3 is highly correlated with one of the temperature variables, but this variable does not have predictive power for this type, which is surprising.

TABLE IV. SIGNIFICANCE OF ESTIMATED PARAMETERS FOR LM.

	Incident type									
Variable	1	2	3	4	5	6	7	8	9	Avg
Intercept	1	1	1	1	1	4	1	1	1	1.33
Trend	1	5	1	4	3	5	5	5	5	3.78
Wind speed	1	5	5	3	5	5	5	5	1	3.89
Temperature	3	4	5	1	2	5	5	5	5	3.89
Rainfall	1	3	5	5	5	5	5	4	1	3.78
Visibility	5	4	5	4	5	5	5	5	5	4.78
Scaling: 1: $p < 0.001$, 2: $p < 0.01$, 3: $p < 0.05$, 4: $p < 0.1$, 5: $p < 1$										

If we look at Table IV in more detail, it stands out that several types have no weather variables with significant predictive power. Opposed to type 3, this is not surprising for type 6 and 7, since their correlations to the weather variables are relatively low as well. On the other hand, types 1 and 9 are well predicted by the amount of wind and rainfall, which is intuitively explainable as well.

Since the wMAPE is higher, we can conclude that the LM is not very good at predicting relatively busy days (compared to predicting average days). However, the fire brigade is, of course, more interested in when they have busy days. They are prepared for average days anyway.

B. Generalized Linear Model

Recall that the GLM model is an expanded version of the linear model, so it could be at least as good. The question is how much value it adds to the linear model. Comparing the significance of the variables in Table V to that of LM in Table IV, we observe that, in general, the single weather

TABLE V. SIGNIFICANCE OF ESTIMATED PARAMETERS FOR GLM.

	Incident type									
Variable	1	2	3	4	5	6	7	8	9	Avg
Intercept	1	2	1	1	1	5	1	2	3	1.89
Trend	1	5	1	4	3	5	5	5	5	3.78
Wind speed	3	5	5	5	5	5	5	5	1	4.33
Temperature	5	5	5	3	2	5	5	5	4	4.33
Rainfall	5	3	5	5	5	5	5	5	1	4.33
Visibility	5	5	4	5	5	5	5	5	5	4.89
Wind*Temp.	5	5	5	5	5	5	5	5	5	5.00
Wind*Rain	3	3	5	5	5	5	5	5	1	4.11
Wind*Visib.	5	3	5	4	5	5	5	5	5	4.67
Temp.*Rain	2	3	5	5	5	5	5	5	1	4.00
Temp.*Visib.	5	5	5	5	5	5	5	5	5	5.00
Rain*Visib.	5	5	5	5	5	5	5	3	5	4.78
Scaling: 1: $p < 0.001$, 2: $p < 0.01$, 3: $p < 0.05$, 4: $p < 0.1$, 5: $p < 1$										

variables have lost some importance in favor of cross-term variables they partition in. Type 1 is an excellent example of this. Here, the temperature had some predictive power in the LM, but now it turns out that it is mainly the *combination* with the amount of rainfall that matters. In addition, also wind speed and rainfall turn out to be less predictive on their own than the LM indicated. It is their cross-term effect that is important. Looking at the average column on the right, we also see that the intercept has lost some importance. Apparently, a bigger part can be modeled by the weather after adding some crossterm variables. Of all weather variables, it is even the case that two cross-term variables have the most predictive power.

Noting the influence of the cross-term variables, we expect that the performance of the GLM is better than that of the LM. If we compute the results for the totals per day, we still see that the wMAPE is somewhat higher than the MAPE, but compared to their equivalents of the LM, they are slightly better (about 2%).

C. Random Forests

Different from the previous models, the RF algorithm does not estimate a parameter for each variable. We, therefore, have to find another measure for the importance of each variable. We will consider the 'RSS-ranking' for this purpose.

In the RF algorithm, in each decision node, the algorithm splits the remaining sample based on a decision rule on the variable that reduces the standard deviation most. In other words, it tries to improve the fit of the model to the training data as much as possible, i.e., the biggest decrease in *residual sum of squares* (RSS) between the fitted model and the observation data in the training set. Hence, we can measure the importance of a variable based on the total decrease in RSS from splitting on this variable. Table VI shows the results of the RSS ranking. As in the previous models, visibility is often the least important variable. However, the biggest difference is that in this case, the temperature is remarkably important.

When we compare the results of RF to the previous models, we see that, in general, RF gives the worst results. However, the effort for running this model is perhaps not in vain. When diving deeper into the results, we discover that the RF has the best wMAPE for type 9, which may be an indication that this algorithm is better in predicting busy days. This is confirmed by the plot of the predictions for type 9 of both GLM and

TABLE VI. IMPORTANCE W.R.T. TOTAL DECREASE IN RSS.



RF in Figure 4. Obviously, the RF algorithm recognizes much better than GLM when the weather conditions are risky and likely to cause many incidents to happen.



Figure 4. Forecasts (in blue) of the number of trucks used for small incidents of type 9, including the upper bound of its 95%-prediction interval (in red).

D. Ensemble model

From the previous discussion, we can conclude that GLM gives the best results when we look at the totals per day, but it is worse in predicting busy days than RF. If we can combine both models in such a way that we capture the good features from both models, then this may improve our forecasts. We will try to do this by applying a form of so-called *ensemble averaging* (EA). In our case, we will take a weighted average of the forecasts of RF and GLM, i.e.,

$$\mathbf{EA} = \gamma \cdot \mathbf{RF} + (1 - \gamma) \cdot \mathbf{GLM},$$

for some constant $\gamma \in [0, 1]$.

We have to determine the optimal value of γ to use in order to get the best results. Since GLM initially gives the best results, and we only need RF to be able to predict the busy days a bit better, we may expect that we have to put more weight on GLM, i.e., that $\gamma < 0.5$. When we vary γ from 0 to 1, both the MAPE = 0.1853 and the wMAPE = 0.1860 take their minimum in $\gamma^* = 0.2$ (which is better than GLM individually; when compared with $\gamma = 0$).

E. Practical implication

After the forecasts are complete, we extract from them the capacity we expect each fire station to need each day. For this, we want to have some certainty that the capacity is satisfying for that day. Different from a confidence interval, which only measures the uncertainty of the forecast, a prediction interval includes, in addition, the variability of the number of incidents in real life. We can, therefore, use the upper bound of the prediction interval to ensure that the predicted capacity will be satisfactory with, for instance, 95% certainty.

The $100(1 - \alpha)$ %-prediction interval for the GLM model $y = X^{\mathsf{T}}\beta + \epsilon$ for a future observation y_0 can be computed as

$$\widehat{y}_0 \pm t_{n-k}^{(1-\alpha/2)} \widehat{\sigma} \sqrt{x_0^{\mathsf{T}} (X^{\mathsf{T}} X)^{-1} x_0 + 1},$$

(see [11]), where \hat{y}_0 is the predicted value for y_0 , $t_{n-k}^{(1-\alpha/2)}$ is the $(1-\alpha/2)$ -quantile of the *t*-distribution with n-k degrees of freedom, n is the number of samples in the training set, and k is the number of variables in the model.

For the RF algorithm, we have N decision trees, which all yield one prediction for each future observation. The variability of these N individual predictions captures the uncertainty of the final prediction (the average of the individuals). In order to capture the variability of the observations, we need again our assumption on the residuals. In this case, we will use this by adding to each of the N individual predictions a random value, drawn from the empirical distribution of the residuals in the training set. Then, the resulting N values include all the variation we need. Their $(\alpha/2)$ - and $(1 - \alpha/2)$ -quantiles together directly form the desired prediction interval.

TABLE VII. CAPACITY NEEDED PER DAY AND FIRE STATION WITH CERTAINTY THIS CAPACITY SUFFICES THAT DAY.

	Avg	cap. nee	eded	% of	Available		
Fire station	90%	95%	99%	90%	95%	99%	cap. 1?
Aalsmeer	0.14	0.17	0.27	0.0%	0.0%	0.0%	No
Amstelveen	0.44	0.53	0.80	0.0%	0.3%	3.3%	No
Anton	0.40	0.48	0.73	0.0%	0.0%	0.3%	No
Diemen	0.12	0.15	0.25	0.0%	0.0%	0.0%	No
Dirk	0.34	0.41	0.64	0.0%	0.0%	0.7%	No
Driemond	0.04	0.05	0.10	0.0%	0.0%	0.0%	Yes
Duivendrecht	0.17	0.20	0.30	0.0%	0.0%	0.0%	No
Hendrik	0.59	0.71	1.07	0.7%	1.7%	67.7%	No
IJsbrand	0.19	0.24	0.38	0.0%	0.0%	0.0%	Yes
Landelijk Noord	0.04	0.06	0.11	0.0%	0.0%	0.0%	Yes
Nico	0.35	0.42	0.64	0.0%	0.0%	0.3%	No
Osdorp	0.42	0.51	0.77	0.0%	0.0%	1.0%	No
Ouderkerk a/d Amstel	0.06	0.08	0.13	0.0%	0.0%	0.0%	Yes
Pieter	0.41	0.50	0.75	0.0%	0.0%	1.7%	Yes
Teunis	0.28	0.34	0.53	0.0%	0.0%	0.0%	No
Uithoorn	0.12	0.15	0.25	0.0%	0.0%	0.0%	No
Victor	0.28	0.34	0.51	0.0%	0.0%	0.0%	No
Willem	0.30	0.36	0.55	0.0%	0.0%	0.0%	No
Zebra	0.23	0.28	0.44	0.0%	0.0%	0.0%	Yes

If we combine all these results, we get Table VII that gives the needed capacity for each fire station. From this, we can conclude that, on an average day, (almost) all fire stations only need a capacity of one truck. Only if we want to be 99% sure that the capacity suffices, we need a capacity of two trucks at station 'Hendrik' on an average day. Then 'Amstelveen' also needs a capacity of two on some days. Moreover, 'Pieter' does not have the required capacity in 1.7% of the days (see in red).

V. CONCLUSIONS AND DISCUSSION

In this paper, we developed a model to create a good forecast on the number of incidents that each fire station in Amsterdam-Amstelland has to handle. Here, special interest went to the influence of several weather conditions and to the issue of dealing with the low number of incidents.

The answer is split into two parts. The forecasts created for the small incidents can be done reasonably well by ensemble averaging (EA). Big incidents can be modeled by an inhomogeneous Poisson process. Concerning the weather, (the combination of) rain and wind on average had the most influence in the linear models and temperature appeared to contain mostly non-linear relations with the number of incidents. As expected beforehand, the visibility had the least predictive power among those four weather variables.

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