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# How Often Are Two Permutations Comparable?

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Boris Pittel The Ohio State University

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#### How often are two permutations comparable?

#### Adam Hammett<sup>1</sup> (joint work with Boris Pittel<sup>2</sup>)

<sup>1</sup>Department of Mathematical Sciences Bethel College

> <sup>2</sup>Department of Mathematics The Ohio State University

> > October 9, 2012

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Comparability of Permutations

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**BAR 4 BA** 

#### Talk Outline



Preliminaries

- Basic Concepts
  - Bruhat Order
  - Weak Order
- Problems Studied



#### Results

- Main Results
- Sketch of Proofs



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Comparability of Permutations

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- (𝔅<sub>n</sub>, ≤) is only a *partially*-ordered set (poset), i.e. it may happen that given π, σ are incomparable.
- Bruhat ordering can be extended to general Coxeter groups, but we studied G<sub>n</sub> only.

**Example.** 3412 > 1324:

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- Despite no restrictions on which inversions we destroy, the total number of inversions strictly decreases each time we destroy one.
- **Problem:** From the definition alone, checking Bruhat comparability is far from algorithmic.
- To get around this, we used two comparability criteria that are algorithmic in nature: the Ehresmann Tableaux and {0, 1}-matrix criteria.

• Discovered by C. Ehresmann in 1934.

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Comparability of Permutations

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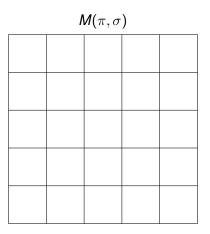
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- Then  $\pi \leq \sigma$  iff the tableau for  $\pi$  is dominated entry-wise by that for  $\sigma$ .

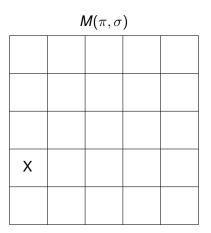
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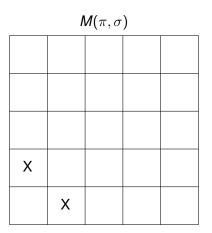
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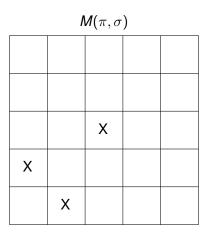
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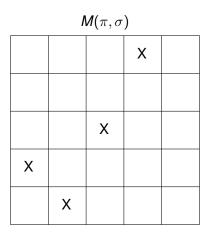
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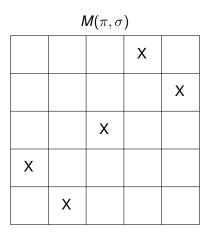
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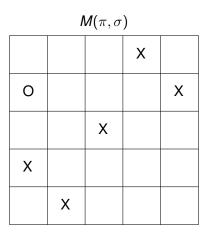
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$M(\pi,\sigma)$							
	0		х				
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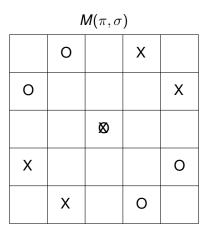
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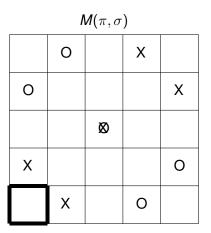
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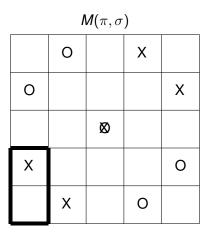
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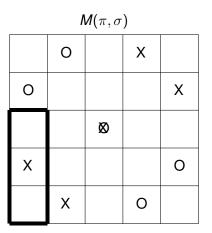
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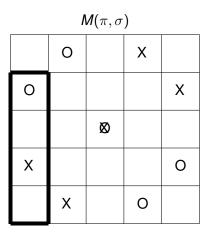
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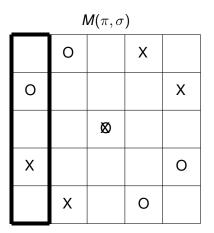
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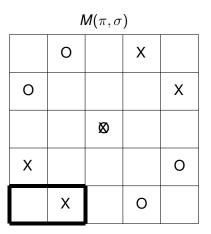
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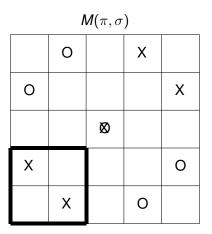
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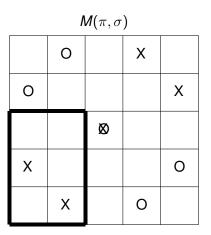
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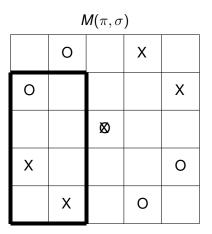
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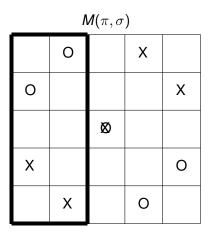
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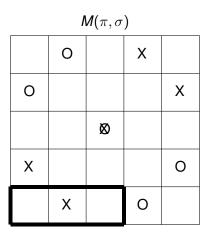
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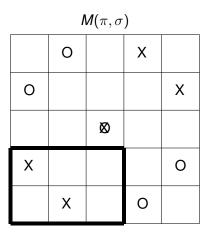
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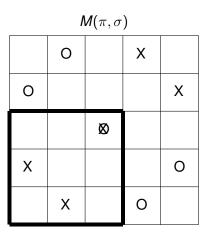
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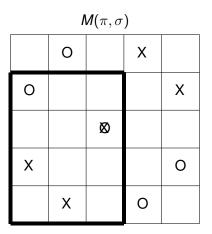
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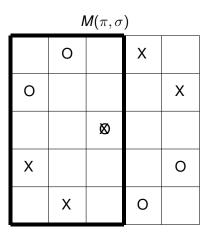
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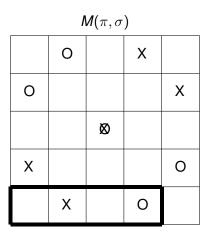
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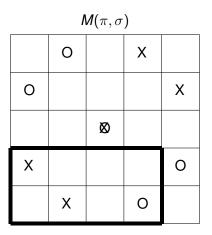
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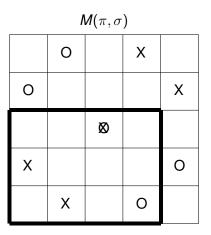
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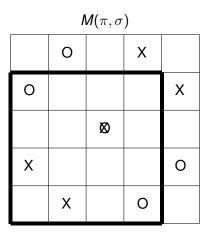
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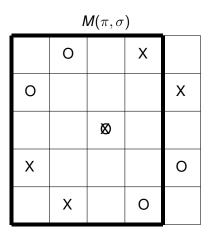
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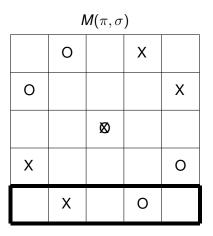
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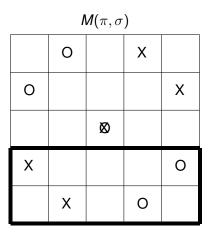
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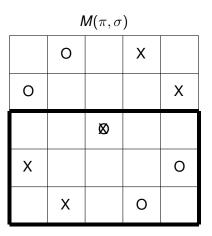
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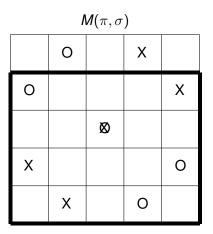
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# The $\{0, 1\}$ -criterion (our primary focus)

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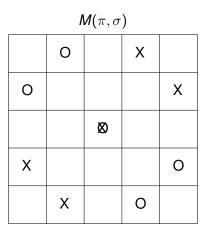
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0				х			
		Ø					
х				0			
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# The $\{0, 1\}$ -criterion (our primary focus)

**Example.** Let  $\pi = 21354$ ,  $\sigma = 45312$ . Get that  $\pi < \sigma$ . **Advantage:** Algorithmic way to check comparability.

$M(\pi,\sigma)$						
	0		Х			
0				х		
		Ø				
х				0		
	Х		0			

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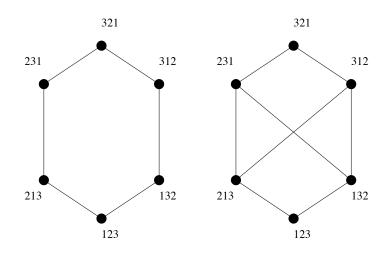
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# The Posets $(\mathfrak{S}_3, \preceq)$ and $(\mathfrak{S}_3, \leq)$



# Equivalent Definition of Weak Order

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Then  $\bigcup_{i \in [n]} \{(j, i) : j \in E_i(\pi)\}$  is the set of non-inversions of  $\pi$ . We have  $\pi \leq \sigma$  iff  $E_i(\pi) \supseteq E_i(\sigma)$  for each  $i \in [n]$ .

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(1)  $(\mathfrak{S}_n, \leq)$  is only *partially*-ordered. So how likely is it that for independent, uniformly random  $\pi, \sigma \in \mathfrak{S}_n$  we have  $\pi \leq \sigma$ ? That is, what are bounds for  $P(\pi \leq \sigma)$ ? (Skandera, MIT, 2004.)

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- (4) As  $(\mathfrak{S}_n, \preceq)$  is a lattice, how likely is it that independent, uniformly random  $\pi_1, \ldots, \pi_r \in \mathfrak{S}_n$  have minimal infimum,  $12 \cdots n$ ?

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  - Pittel studied the analogous problems for the poset of integer partitions under dominance order, and for the poset of set partitions ordered by refinement.

### Main Results

# **Bruhat Order Results**

### Theorem

Let  $\pi_1, \ldots, \pi_r \in \mathfrak{S}_n$  be independent and uniformly random. Then there are uniform constants  $c_1 = c_1(\epsilon), c_2 > 0$  such that

$$c_1\left(\frac{1}{r!}-\epsilon\right)^n \leq P(\pi_1 \leq \cdots \leq \pi_r) \leq c_2 n^{-r(r-1)}, \quad \forall \epsilon > 0.$$

Equivalently, there are at least  $(n!)^r c_1(1/r! - \epsilon)^n$  and at most  $(n!)^r c_2 n^{-r(r-1)}$  length *r* chains in Bruhat order. In the case r = 2, there is a uniform constant c > 0 such that

$$P(\pi_1 \leq \pi_2) \geq c(0.708)^n.$$

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• We will focus on the proof of the r = 2 upper bound.

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# Weak Order Comparability Results

### Theorem

Let  $\pi, \sigma \in \mathfrak{S}_n$  be independent, uniformly random, and write  $P_n^* := P(\pi \leq \sigma)$ . Then, as a function of n,  $P_n^*$  is submultiplicative, i.e.  $P_{n_1+n_2}^* \leq P_{n_1}^* P_{n_2}^*$ . So (Fekete lemma) there exists  $\rho = \lim_n (P_n^*)^{1/n} = \inf_k (P_k^*)^{1/k}$ . Furthermore, there exists an absolute constant c > 0 such that

$$\prod_{i=1}^{n} H(i)/i \le P_{n}^{*} \le c(0.362)^{n};$$

here  $H(i) = \sum_{i=1}^{i} 1/j$ . Consequently  $\rho \leq 0.362$ .

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• Note that in any case (Bruhat or weak ordering) we have  $P(\bullet) \rightarrow 0, n \rightarrow \infty$ .

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# Weak Order Lattice-property Results

### Theorem

Write  $P_{n,r} := P(\inf\{\pi_1, \dots, \pi_r\} = 12 \cdots n)$ . Then, as a function of n,  $P_{n,r}$  is submultiplicative, and

$$\lim_{n\to\infty}(P_{n,r})^{1/n}=1/z^*;$$

here,  $z^* = z^*(r) \in (1, 2)$  is the unique (positive) root of the equation  $\sum_{j \ge 0} (-1)^j z^j / (j!)^r = 0$  within the disk  $|z| \le 2$ .

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• In the case r = 2, we have  $1/z^* \approx 0.69$ . Note that, for r fixed,  $P_{n,r} \rightarrow 0$  exponentially fast as  $n \rightarrow \infty$ .

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Results

Sketch of Proofs

Toward the Proof of the Bruhat Order Upper Bound (r=2)

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- We need to select a subset of conditions necessary for  $\pi \leq \sigma$  that are sufficiently simple, so that we can compute (estimate) the number of these pairs.
- On the other hand, these conditions need to stringent enough so that they collectively have probability o(1).

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## Toward the Proof of the Bruhat Order Upper Bound

 First Advance: the Ehresmann Criterion implies that for each k ≤ n,

$$\{\pi \le \sigma\} \subseteq \left\{ \sum_{i=1}^{j} \pi(i) \le \sum_{i=1}^{j} \sigma(i), \, \forall j \le k \right\}.$$
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- So long as  $k = o(n^{1/2})$ , the first k entries of a random permutation are asymptotically independent and uniform on [n].
- So letting  $k \to \infty$  "slowly" with *n*, we obtain from (1)

$$P(\pi \leq \sigma) = O\left(n^{-1/2}\right)$$

by using a certain connection with a random walk on the real line (Feller, "Intro. to Prob. Theory, Vol. II").

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$$\pi(i_1) \leq \sigma(j_1), \ldots, \pi(i_k) \leq \sigma(j_k).$$

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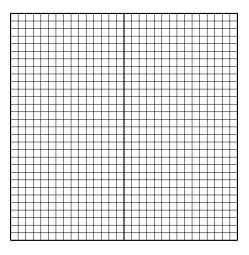
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• This is the same as "reading-off" rows of the first *k* columns of  $M(\pi, \sigma)$ , bottom to top, with the # X's (for  $\pi$ ) always more than the # O's (for  $\sigma$ ) at any intermediate point.

Sketch of Proofs

## Equivalence of Ehresmann and {0, 1}-criteria



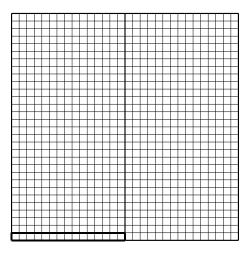
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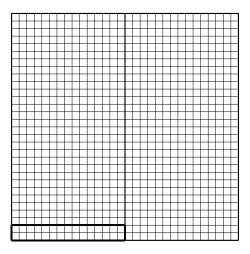
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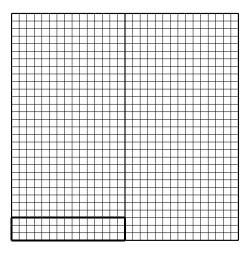
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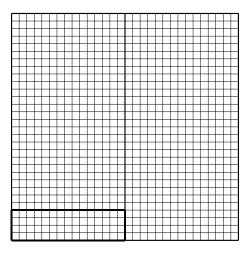
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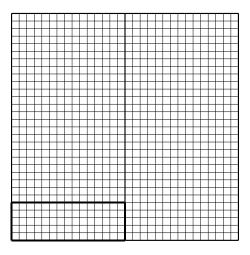
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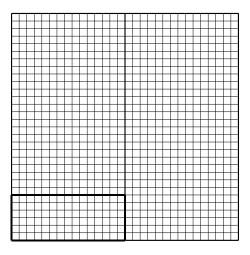
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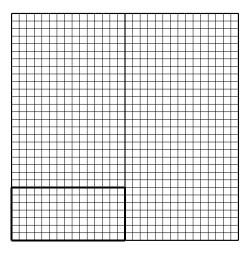
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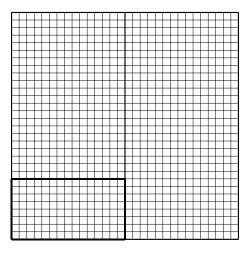
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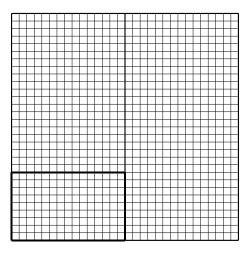
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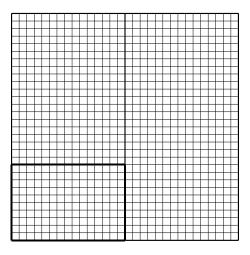
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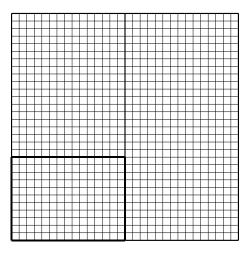
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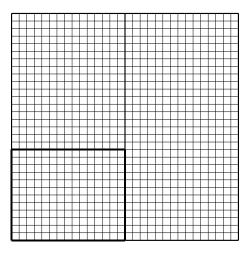
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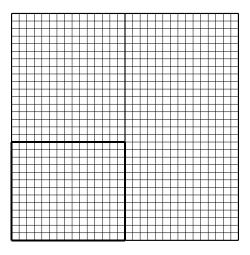
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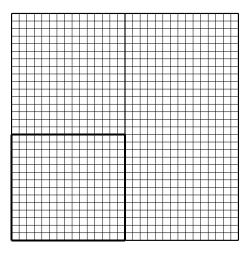
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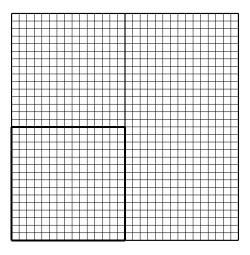
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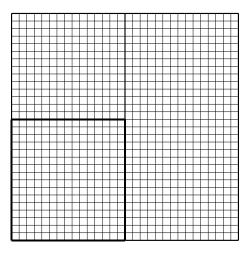
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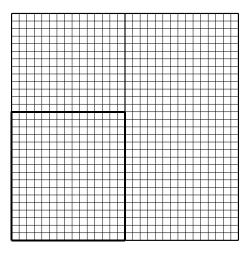
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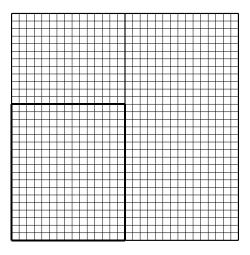
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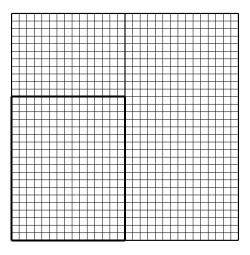
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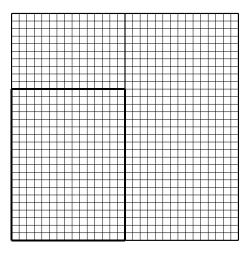
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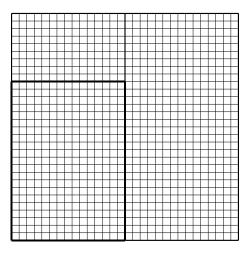
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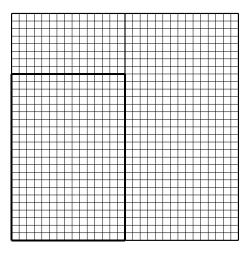
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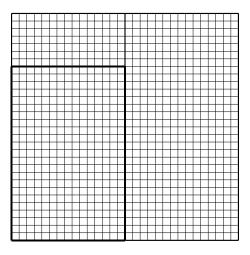
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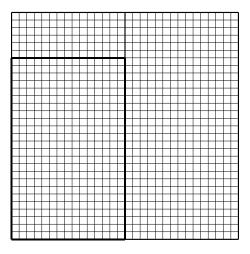
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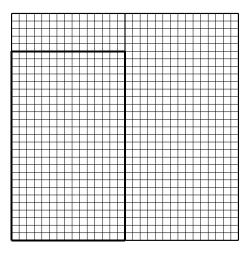
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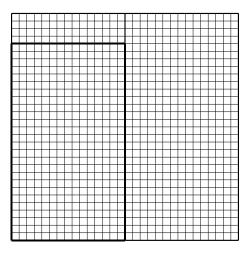
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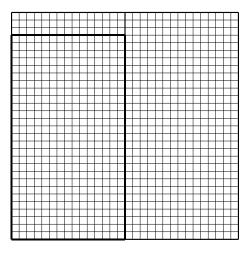
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Sketch of Proofs

# Equivalence of Ehresmann and {0, 1}-criteria



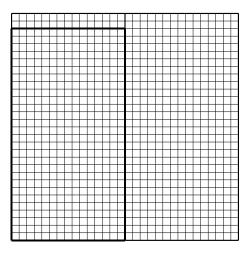
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Sketch of Proofs

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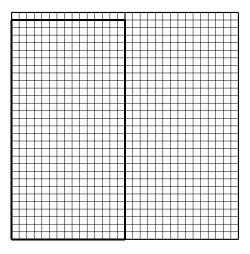
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Sketch of Proofs

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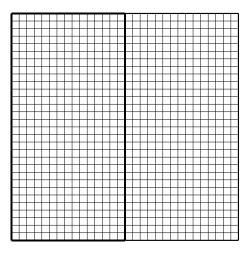
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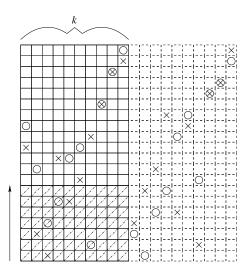


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# Equivalence of Ehresmann and {0, 1}-criteria



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- Other  $2(k \ell)$  rows need to be split between  $\pi$  (for X's) and  $\sigma$  (for O's) according to the ballot condition.

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- $\ell := #$  of rows with both an X and O.
- Other  $2(k \ell)$  rows need to be split between  $\pi$  (for X's) and  $\sigma$  (for O's) according to the ballot condition.
- The number of ways to do this, for given  $\ell$ , is

$$\binom{n}{\ell}\binom{n-\ell}{2(k-\ell)}\frac{1}{k-\ell+1}\binom{2(k-\ell)}{k-\ell},$$

the last two factors coming from the classic Ballot Theorem.

# Statement of the Ballot Theorem

### Theorem

Candidate A receives a votes, B gets b votes, a > b. Then the number of ballot tallies (counted 1 vote at a time) such that A is always strictly ahead of B equals

$$\frac{a-b}{a+b}\binom{a+b}{a}.$$

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• Equ., starting at (0, 0), we can make a rightward unit move each time *A* gets a vote, and an upward unit move each time *B* gets a vote. Then this theorem counts the number of lattice paths with these moves, joining the points (0, 0) - (a, b), that never *touch* the diagonal y = x.

## Ballot Theorem cont.

• In our case, we are *allowed* to touch the diagonal, as "ties" in the cumulative counts are permitted.

#### Sketch of Proofs

## Ballot Theorem cont.

- In our case, we are *allowed* to touch the diagonal, as "ties" in the cumulative counts are permitted.
- To compensate for this, we "shift" the diagonal left 1 unit, and the Ballot Theorem count changes to, for a ≥ b,

$$\frac{a+1-b}{a+1+b}\binom{a+1+b}{a+1} = \frac{a-b+1}{a+1}\binom{a+b}{a}.$$

### Sketch of Proofs

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• For us,  $a = b = k - \ell$ , and this delivers the count

$$\binom{n}{\ell}\binom{n-\ell}{2(k-\ell)}\frac{1}{k-\ell+1}\binom{2(k-\ell)}{k-\ell}$$

we claimed for the total number of admissible row selections for  $\pi$  (to contain X's) and for  $\sigma$  (to contain O's) with overlap size  $\ell$ .

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 To complete the construction of pairs (π, σ) ∈ A<sub>k</sub>, we need to decide where to *put* the X's and O's in these chosen rows, and also place the remaining n − k X's and n − k O's somewhere in the remaining rows/columns. The total number of ways to do this is

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• Putting these pieces together, and summing over all  $\ell \leq k$ , we get

$$P(\pi \le \sigma) \le P(A_k) = \sum_{\ell \le k} \frac{\binom{n}{\ell} \binom{n-\ell}{2(k-\ell)} \binom{2(k-\ell)}{k-\ell} (k!)^2 (n-k)!^2}{(n!)^2 (k-\ell+1)}$$
$$= \frac{n+1}{(n-k+1)(k+1)} \sum_{\ell \le k} \frac{\binom{k}{\ell} \binom{n+1-k}{k+1-\ell}}{\binom{n+1}{k+1}}$$
$$= \frac{n+1}{(n-k+1)(k+1)} = O\left(n^{-1}\right).$$

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For fixed k and n→∞, P(A<sub>k</sub>) ~ (k + 1)<sup>-1</sup>, which is in accordance with our intuition.

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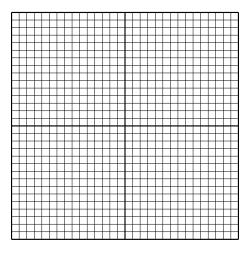
 To obtain the final result, P(π ≤ σ) = O(1/n<sup>2</sup>), how do we take into account an even larger subset of conditions?

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- Notice that we did not pay attention to the conditions in the last *n* - *k* columns. So we need to incorporate them somehow, while still preserving our ability to enumerate the resulting pairs of permutations.

- To obtain the final result,  $P(\pi \le \sigma) = O(1/n^2)$ , how do we take into account an even larger subset of conditions?
- Notice that we did not pay attention to the conditions in the last *n* - *k* columns. So we need to incorporate them somehow, while still preserving our ability to enumerate the resulting pairs of permutations.
- With the ballot-like conditions we just encountered driving our intuition, we arrive at the following picture:

Sketch of Proofs

# Finding a Necessary Condition for $\pi \leq \sigma$



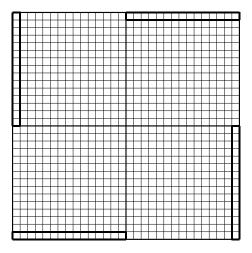
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Sketch of Proofs

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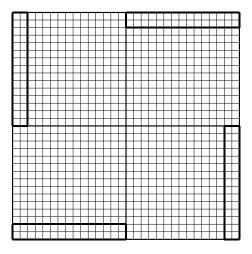
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Sketch of Proofs

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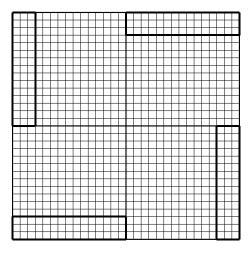


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Sketch of Proofs

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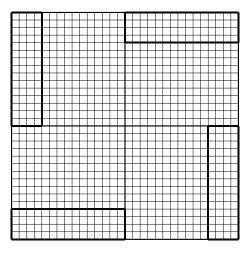
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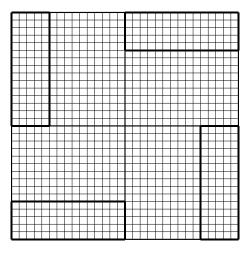
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Sketch of Proofs

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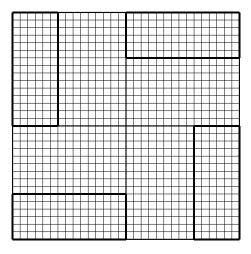
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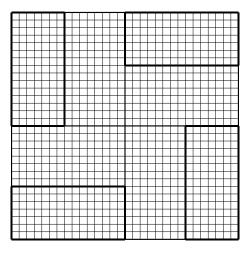
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Sketch of Proofs

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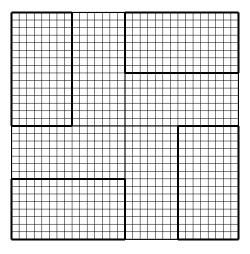
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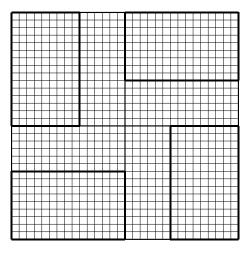
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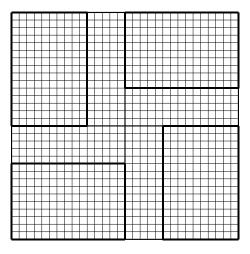
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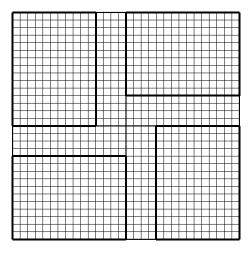
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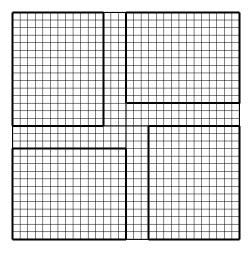
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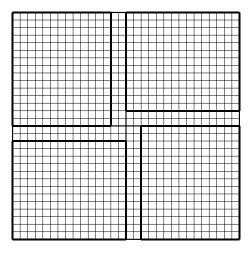
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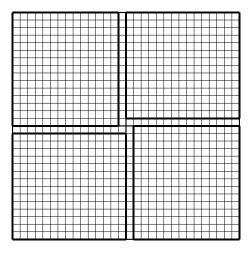
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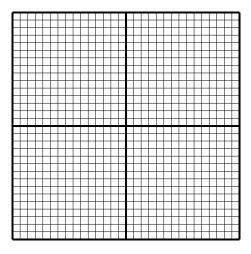
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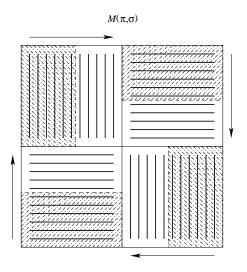


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# Finding a Necessary Condition for $\pi \leq \sigma$



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• In a manner similar to the O(1/n) proof, we obtain

$$\begin{split} P(\pi \leq \sigma) \\ \leq & \sum_{m_1 \geq m_2} \frac{(m_1 - m_2 + 1)^4 (n/2 + 1)^4}{(m_1 + 1)^4 (n/2 - m_2 + 1)^4} \binom{n/2}{m_1}^4 \binom{n/2}{m_2}^4 \\ & \times \frac{m_1!^2 (n/2 - m_1)!^2 m_2!^2 (n/2 - m_2)!^2}{n!^2} \\ = & \sum_{m_1 \geq m_2} \frac{(m_1 - m_2 + 1)^4 (n/2 + 1)^4}{(m_1 + 1)^4 (n/2 - m_2 + 1)^4} \prod_{i=1}^2 \frac{\binom{n/2}{m_i} \binom{n/2}{n/2 - m_i}}{\binom{n}{n/2}}. \end{split}$$

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Extending this last sum over all *m*<sub>1</sub>, *m*<sub>2</sub> (not just *m*<sub>1</sub> ≥ *m*<sub>2</sub>), we see that the extended sum equals

$$\mathsf{E}\left[\frac{(M_1-M_2+1)^4(n/2+1)^4}{(M_1+1)^4(n/2-M_2+1)^4}\right],$$

so that this expectation bounds our probability  $P(\pi \leq \sigma)$  from above.

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so that this expectation bounds our probability  $P(\pi \leq \sigma)$  from above.

• Here,  $M_1$ ,  $M_2$  are independent copies of the Hypergeometric random variable with parameters n/2, n/2, n/2. So  $M_i$  is equal in distribution to the number of red balls in a uniformly random sample of size n/2 from a bin containing n/2 red and n/2 white balls.

To finish our bound on P(π ≤ σ), it remains to estimate this expectation from above.

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- The  $M_i$  are sharply concentrated around their mean, n/4, with exponentially high probability. Further, the difference  $|M_1 M_2 + 1|$  has expectation of order  $\sqrt{n}$  at most.

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- The  $M_i$  are sharply concentrated around their mean, n/4, with exponentially high probability. Further, the difference  $|M_1 M_2 + 1|$  has expectation of order  $\sqrt{n}$  at most.
- So, roughly speaking, we conclude that

$$egin{split} \mathcal{P}(\pi \leq \sigma) \leq \mathcal{E}\left[rac{(M_1-M_2+1)^4(n/2+1)^4}{(M_1+1)^4(n/2-M_2+1)^4}
ight] \ &= \mathcal{O}\left(rac{(\sqrt{n})^4 \cdot n^4}{n^4 \cdot n^4}
ight) = \mathcal{O}\left(n^{-2}
ight). \end{split}$$

# Conjectures

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# Conjectures

• Write 
$$P_n = P(\pi \leq \sigma), P_n^* = P(\pi \leq \sigma).$$

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# Conjectures

- Write  $P_n = P(\pi \leq \sigma), P_n^* = P(\pi \leq \sigma).$
- (1) There is  $\delta \in [0.5, 1]$  and C > 0 such that  $P_n \sim Cn^{-(2+\delta)}$ .

#### Sketch of Proofs

# Conjectures

- Write  $P_n = P(\pi \leq \sigma), P_n^* = P(\pi \leq \sigma).$
- (1) There is  $\delta \in [0.5, 1]$  and C > 0 such that  $P_n \sim Cn^{-(2+\delta)}$ .
- (2) There is  $\rho \in [0.3, 1/3]$  and C > 0 such that  $P_n^* \sim C \rho^n$ . Here

$$\rho = \lim_{n \to \infty} \sqrt[n]{P_n^*}.$$

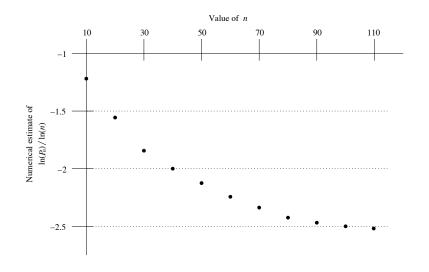
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# **Bruhat Order Numerics**

п	R <sub>n</sub>	Estimate of $P_n \approx \frac{R_n}{10^9}$	Estimate of $\ln(P_n)/\ln n$
10	61589126	0.0615891	-1.21049
30	1892634	0.0018926	-1.84340
50	233915	0.0002339	-2.13714
70	50468	0.0000504	-2.32886
90	14686	0.0000146	-2.47313
110	5174	0.0000051	-2.58949

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# **Bruhat Order Numerics**



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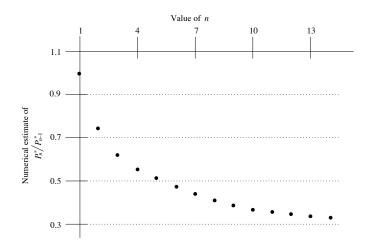
#### Weak Order Numerics

n	$R_n^*$	Estimate of $P_n^* \approx \frac{R_n^*}{10^9}$	Estimate of $P_n^*/P_{n-1}^*$
10	1538639	0.0015386	0.368718
11	541488	0.0005414	0.351926
12	184273	0.0001842	0.340308
13	59917	0.0000599	0.325153
14	18721	0.0000187	0.312448
15	5714	0.0000057	0.305218
16	1724	0.0000017	0.301715

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#### Weak Order Numerics



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Any permutation matrix is also an alternating sign matrix.

The set of Monotone Triangles of order n,  $\mathfrak{M}_n$ :

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$$\begin{pmatrix} 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & -1 & 0 & 0 & 1 \\ 0 & 1 & -1 & 1 & 0 \\ 0 & 0 & 1 & 0 & 0 \end{pmatrix} \longleftrightarrow \begin{pmatrix} 0 & 1 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 & 1 \\ 1 & 1 & 0 & 1 & 1 \\ 1 & 1 & 1 & 1 & 1 \end{pmatrix}$$

H. and Pittel (BC, OSU)

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$$\begin{array}{c} 2 \\ 2 & 3 \\ \longleftrightarrow & 1 & 3 & 5 \\ 1 & 2 & 4 & 5 \\ 1 & 2 & 3 & 4 & 5 \end{array}$$

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- What about comparability probability for this lattice? Recent work has only focused on enumeration of these objects (Zeilberger, Kuperberg).
- What about the size of the largest *anti*-chain in weak order? This is closed for Bruhat order (it has the Sperner property; Engel, "Sperner Theory").

## For Further Reading



A. Hammett, B. Pittel. How often are two permutations comparable? *Trans. of the Amer. Math. Soc.*, 2009.

H. and Pittel (BC, OSU)

Comparability of Permutations

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