A distance for probability spaces, and long-term values in Markov Decision Processes and Repeated Games

Jérôme Renault^{*}, Xavier Venel[†]

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Abstract

Given a compact subset X of a normed vector space, we study the pseudo-metric on Borel probabilities over X given by $d_*(u, v) = \sup_{f \in D_1} |u(f) - v(f)|$, where D_1 is the set of functions satisfying: $\forall x, y \in X, \forall a, b \geq 0$, $af(x) - bf(y) \leq ||ax - by||$. The particular case where X is a simplex endowed with the L^1 -norm is particularily interesting: in this case d_* can be characterized as the largest distance on the probabilities with finite support over X which makes all disintegrations non expansive. Moreover we obtain a Kantorovich-Rubinstein type duality formula for $d_*(u, v)$ involving couples of measures (α, β) over $X \times X$ such that the first marginal of α is u and the second marginal of β is v.

In the second part of the paper, we study several kinds of Markov Decision Processes, Gambling Houses and 2-player zero-sum Repeated Games. It includes in particular all partial observation MDPs with finitely many states, and repeated games with an informed controller and finite sets of states and actions. In each case, the underlying state space is compact and the transitions can be shown to be non expansive for the distance d_* (or in some case for the Kantorovich-Rubinstein distance). This allows us to prove the existence of, and to characterize via the introduction of appropriate invariant measures, a very strong notion of limit value called the general uniform value. The decision-maker is able to play well independently of the time horizon, and regarding evaluations for payoffs we consider not only the Cesàro means when the number of stages is large, but any evaluation function θ over stages when the total variation, or impatience, $TV(\theta) = \sum_{t>1} |\theta_{t+1} - \theta_t|$ is small enough.

Keywords: Markov decision processes, gambling houses, POMDPs, repeated games, distance for belief spaces, Kantorovich-Rubinstein duality, Wasserstein metric, limit value, uniform value, general values, characterization of the value.

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^{*}TSE (GREMAQ, Université Toulouse 1 Capitole), 21 allée de Brienne, 31000 Toulouse, France. E-mail: jerome.renault@ut-capitole.fr

[†]The School of Mathematical Sciences, Tel Aviv University, Tel Aviv 69978, Israel. E-mail: xavier.venel@gmail.com

1 Introduction

Given a compact metric space X, we denote by $\Delta(X)$ the set of Borel probabilities over X. The Kantorovich-Rubinstein metric over $\Delta(X)$ can be defined by $d_{KR}(u,v) = \sup_{f \in E_1} |u(f) - v(f)|$, where E_1 is the set of 1-Lipschitz functions, and has been extensively studied in the literature. We assume here that X lies in a normed vector space, and consider the pseudo-metric over $\Delta(X)$ given by: $d_*(u,v) = \sup_{f \in D_1} |u(f) - v(f)|$, where D_1 is the set of 1-Lipschitz functions satisfying: $\forall x, y \in X, \forall a, b \geq 0, af(x) - bf(y) \leq ||ax - by||$. We will prove in Theorem 2.7 a first duality theorem:

 $d_*(u,v) = \inf_{\gamma \in \mathcal{M}_3(u,v)} \int_{X^2 \times [0,1]^2} \|\lambda x - \mu y\| d\gamma(x,y,\lambda,\mu),$ where $\mathcal{M}_3(u,v)$ contains all finite positive measures on $X^2 \times [0,1]^2$ such that for each f continuous,

$$\int_{X^2\times[0,1]^2}\lambda f(x)d\gamma(x,y,\lambda,\mu)=u(f)$$

and

$$\int_{X^2 \times [0,1]^2} \mu f(y) d\gamma(x, y, \lambda, \mu) = v(f).$$

We then concentrate on the case where $X = \Delta(K)$ is a simplex (K is a finite set) endowed with the L^1 -norm, and we denote by $\Delta_f(X)$ the set of probabilities with finite support over X. In this case d_* is a metric metrizing the weak* topology on $\Delta(X)$. For any finite set S, the disintegration mapping ψ_S , defined on $\Delta(K \times S)$ by $\psi_S(\pi) = \sum_{s \in S} \pi(s) \delta_{p(s)}$ where for each s, p(s) is the posterior on K given s, is now 1-Lipschitz from $(\Delta(K \times S), \|.\|_1)$ to $(\Delta_f(X), d_*)$. This seems a desirable property (not shared by the Kantorovich-Rubinstein metric), since π contains some information on s, and certainly more information than $\psi_S(\pi)$. Moreover, we will show that d_* is the largest distance on $\Delta_f(X)$ having this property. Finally, we prove in Theorem 2.19 a duality formula for probabilities u and v with respective finite supports U and V over X:

$$d_*(u,v) = \min_{(\alpha,\beta)\in\mathcal{M}_4(u,v)} \sum_{(x,y)\in U\times V} \|x\alpha(x,y) - y\beta(x,y)\|,$$

where $\mathcal{M}_4(u, v)$ is the set of couples (α, β) of positive measures over $U \times V$ such that the first marginal of α is u and the second marginal of β is v.

In the rest of the paper we use the above results to prove the existence of a strong notion of limit value in Gambling Houses, Markov Decision Processes and Repeated Games.

The standard model of Markov Decision Processes (or Controlled Markov chain) was introduced by Bellman [6] in the 1950s and has been extensively studied since then. In this model, a decision-maker perfectly observes at the beginning of every stage what is the current state, and chooses an action accordingly. The current state and the selected action determine a stage payoff and the law of the next state. Regarding long-term values, it is standard to look at the problem with discounted payoffs (the stream of payoffs is evaluated with a discount factor, typically small), and at the problem with finitely many stages (the stream of payoffs is evaluated by the average of the payoffs of the first T stages, with T large). When the sets of states and actions are both finite, Blackwell [9] proved the existence of a strategy which is exactly optimal for all discount factors close to 0, that is when the decision-maker is sufficiently patient. And in particular the discounted values, when the discount goes to zero, and the T-stage values, when T goes to infinity, converge to the same limit value.

This model was later generalized to MDPs with Partial Observation (POMDP), a good reference being Arapostathis *et al.* [1]. The decision-maker no longer observes the current state, but at the beginning of each stage he receives a signal which depends on the previous and current states and his previous action. In order to study such problem a natural approach is to go back to the standard model of MDPs with full observation on the new state, with new state space the space of beliefs on the original state (see Astrom, K.J. [3], Sawaragi and Yoshikawa [39] and Rhenius [35]). For standard optimality criteria such as the *T*-stage values or the discounted values, these two problems are equivalent and the question of the existence of the limit value is the same.

Part of the literature assumes some ergodicity conditions to get existence and characterization of the limit value. In this setup one typically searches for a solution of the "Average Cost Optimality Criterion", or ACOE, in order to find the limit value of the MDP, for example as in Runggaldier and Stettner [38] or as in Borkar [10], [12]. An introduction to the ACOE in the framework of MDP and the reduction of POMDP can be found in Hernández-Lerma [19]. From another point of view, if we know that the limit value exists, the ACOE may be used to obtain a characterization of the value. For standard MDPs with finite states and actions, Denardo and Fox [14] proved that the limit value is the solution of a linear programming problem deduced from the ACOE. Moreover by standard linear programming results, it is also the solution of a dual problem described by Hordjik and Kallenberg [20] (see Remark 3.22). This dual problem focuses on the maximal payoff that the decision-maker can guarantee on invariant measures. This approach was extended to different criteria (see Kallenberg [22]) and to a convex analytic approach by Borkar (for references see Borkar [11] in order to study problems with a countable state space and a compact action space.

We will avoid here any ergodicity condition, consequently the limit value will depend on the initial state and we will obtain, for each $\varepsilon > 0$, the existence of strategies that are ε -optimal in the long run (see example 3.15). We will also avoid conditions on the actions set and concentrate on the state space, the transitions and the payoffs. Given an initial POMDP on a finite space K, we will follow the natural approach and introduce a MDP on $X = \Delta(K)$ but instead of assuming some ergodicity on the control we will use the structure of $\Delta(K)$ and our new metric on $Z = \Delta_f(\Delta(K))$. We extend and relax the MDP on Z, and using the metric d_* , we are able to obtain *non-expansive* affine transitions. So the transitions become 1-Lipschitz on the totally bounded space Z, and this is a main point of our approach.

We are able to strengthen the notion of limit value and deal with general evaluations. Given a probability distribution θ on positive integers, a sequence of payoffs $(g_t)_{t\geq 1}$ is evaluated by $\sum_{t\geq 1} \theta_t g_t$. In a MDP or a POMDP, the θ -value is then defined as the supremum expected payoff that the player can guarantee with this evaluation. We will consider the case of a "patient" decision-maker, in the sense that the total variation $TV(\theta) = \sum_{t \ge 1} |\theta_{t+1} - \theta_t|$ will be small, or go to zero. Most of the literature focuses on the T-stage values, when T goes to $+\infty$, and on the λ -discounted values, when λ converges to 0. When there is no control (the set of actions is a singleton), the relation between the two is directly linked to a Hardy-Littlewood theorem (see Filar and Sznajder, [17]): one of the limit exists if and only if the other exists, and in case of existence the limits are the same. Lehrer and Sorin [23] proved that this result extends to the MDP provided we consider uniform convergence of the value functions. Another approach focuses on the existence of a strategy which is ε -optimal in any long enough finite game or for any discount factor small enough. When such a strategy exists for each ε , the MDP is said to have a uniform value. For MDP with finitely many states and actions, Blackwell's result [9] also implies the existence of the uniform value. In POMDPs, Rosenberg *et al.* [36] proved the existence of the uniform value when the sets of states, actions and signals are finite, and Renault [32] removed the finiteness assumption on signals and actions. In the present paper, we will consider an extremely strong notion of limit value called the "general uniform value", encompassing the limit value as well as the uniform value, not only for discounted payoffs or average payoffs but for evaluations θ with $TV(\theta)$ small (see for instance Definitions 3.5 and 3.6).

The extension of standard MDPs to 2-player stochastic games is due to Shapley [40] (and was even published before Bellman, [6]): the state variable is now simultaneously controlled by 2 players having opposite interests. Whenever states and actions are finite, the existence of the limit value is due to Bewley and Kohlberg [7]. A few years later, Mertens and Neyman [26] proved in this setup the existence of the uniform value. Naturally the model of stochastic games also generalizes to partial information (see Mertens [28]), but the existence of possible private information for the different players implies a very complex structure on the auxiliary state space. Following Harsanyi [18], Mertens and Zamir [27] introduced the universal belief space which synthesizes all the information for both players in a general repeated game: their beliefs about the state, their beliefs about the beliefs of the other player, etc... So far, the results of the literature always concern some subclasses of games where we can explicitly write the auxiliary game in a "small" tractable set. A lot of work has especially been done about games with one fully informed player and one player with partial information, and we will only consider such games here. In the simplest model introduced by Aumann and Maschler (see reference from 1995), a state is initially chosen and remains fixed for the rest of the game. Renault [31] extended the analysis to a general underlying Markov chain on the state space (see also Neyman, [30]). Renault [33] proved the existence of the uniform value when the informed player can additionally control the evolution of the state variable (see also Rosenberg *et al.* [37]).

In this paper we first consider Gambling Houses defined by $\Gamma = (X, F, r)$, where X is the state space, $r: X \to [0, 1]$ is the running payoff and $F: X \rightrightarrows \Delta_f(X)$ is the transition multifunction. Given an initial state x_0 in X, a decisionmaker, or player, has to choose u_1 in $F(x_0)$, then x_1 is selected according to u_1 and there is a payoff $r(x_1)$, etc... We show in Theorem 3.10 that if X is metric compact, r is continuous and F is non expansive with respect to the Kantorovich-Rubinstein distance, then the problem has a general uniform value v^* characterized by:

$$\forall x \in X, \ v^*(x) = \inf \left\{ w(x), w : \Delta(X) \to [0, 1] \text{ affine continuous s.t.} \\ (1) \ \forall y \in X, w(y) \ge \sup_{u \in F(y)} w(u) \text{ and } (2) \ \forall u \in R, w(u) \ge r(u) \right\}.$$

where R is a suitably defined set of invariant measures for the Gambling House Γ (see Definition 3.9).

We then introduce explicitly actions and study standard MDPs Ψ given by a set of states X, a non empty set of actions A, a mapping $q: X \times A \to \Delta_f(X)$ and a payoff function $g: X \times A \to [0, 1]$. At each stage, the player learns the current state x and chooses an action a. He then receives the payoff g(x, a), a new state is drawn accordingly to q(x, a) and the game proceeds to the next stage. We assume in Theorem 3.20 that X is a compact subset of a simplex $\Delta(K)$, and moreover that $\forall x \in X, \forall y \in X, \forall a \in A, \forall f \in D_1, \forall a \ge 0, \forall \beta \ge 0,$ $|\alpha f(q(x, a)) - \beta f(q(y, a))| \le ||\alpha x - \beta y||_1$ and $|\alpha g(x, a) - \beta g(y, a)| \le ||\alpha x - \beta y||_1$. Then we prove that Ψ has a general uniform value v^* characterized by: for all xin X,

$$\begin{aligned} v^*(x) &= \inf \quad \left\{ w(x), w : \Delta(X) \to [0, 1] \text{ affine continuous } s.t. \\ (1) \; \forall x' \in X, w(x') \geq \sup_{a \in A} w(q(x', a)) \text{ and } (2) \; \forall (u, y) \in RR, w(u) \geq y \; \right\}. \end{aligned}$$

where RR is a suitably defined set of *invariant couples* for the MDP Ψ (see Definition 3.19). This result can in particular be applied when X is finite and provides an alternative formulation for the limit value in the simplest case (see Remark 3.22). To prove Theorem 3.20, we use the properties of the metric d_* introduced in the first part of the paper.

Finally, we are able to apply the above result to any POMDP with finite set of states (without assumptions on the set of actions), and to any repeated game with an informed controller with finitely many states and actions.

Section 2 of this manuscript is dedicated to the study of the (pseudo)-metric d_* . In section 3, we study long-term values for compact non expansive Gambling Houses and MDPs. Finally section 4 contains the applications to Partial Observation MDP and repeated games with an informed controller.

2 A distance for belief spaces

2.1 A pseudo-distance for probabilities on a compact subset of a normed vector space

We fix a compact subset X of a real normed vector space V. We denote by $E = \mathcal{C}(X)$ the set of continuous functions from X to the reals, and by E_1 the set of 1-Lipschitz functions in E. We denote by $\Delta(X)$ the set of Borel probability measures on X, and for each x in X we write δ_x for the Dirac probability measure on x. It is well known that $\Delta(X)$ is a compact set for the weak-* topology, and this topology can be metrizable by the (Wasserstein) Kantorovich-Rubinstein distance:

$$\forall u, v \in \Delta(X), \ d_{KR}(u, v) = \sup_{f \in E_1} u(f) - v(f).$$

We will introduce a pseudo-distance on $\Delta(X)$, which is not greater than d_{KR} and in some cases also metrizes the weak-* topology. We start with several definitions, which will turn out to be equivalent. Let u and v be in $\Delta(X)$.

Definition 2.1.

$$d_1(u, v) = \sup_{f \in D_1} u(f) - v(f),$$

where $D_1 = \{ f \in E, \forall x, y \in X, \forall a, b \ge 0, af(x) - bf(y) \le ||ax - by|| \}.$

Note that any linear functional in V' with norm 1 induces an element of D_1 . If f is in D_1 then -f is also in D_1 , so $d_1(u, v) = \sup_{f \in D_1} |u(f) - v(f)|$ and d_1 is a pseudo-distance on $\Delta(X)$. We also have $D_1 \subset E_1$, so that $d_1(u, v) \leq d_{KR}(u, v)$ and the supremum in the definition of $d_1(u, v)$ is achieved. $(x \mapsto ||x||)$ belongs to D_1 , consequently for x and y in X we have $d_1(\delta_x, \delta_y) = ||x - y||$.

Notice that $D_1 = \{f \in E, \forall x, y \in X, \forall a, b \ge 0, |af(x) - bf(y)| \le ||ax - by||\}$. If f, g are in D_1 then $\sup\{f, g\}$ and $\inf\{f, g\}$ also are, and D_1 is a convex lattice with greatest element $(x \mapsto ||x||)$ and smallest element $(x \mapsto -||x||)$. If $0 \in X$, then all f in D_1 satisfy f(0) = 0.

Example 2.2. First consider the particular case where X = [0, 1] endowed with the usual norm. Then all f in D_1 are linear. As a consequence, $d_1(u, v) = 0$ for

 $u = 1/2 \,\delta_0 + 1/2 \,\delta_1$ and $v = \delta_{1/2}$. We do not have the separation property and d_1 is not a distance in this case¹.

Let us modify the example. X now is the set of probability distributions over 2 elements, viewed as $X = \{(x, 1 - x), x \in [0, 1]\}$. We use the norm $\|.\|_1$ to measure the distance between (x, 1 - x) and (y, 1 - y), so that $V = \mathbb{R}^2$ is endowed with $\|(x_1, x_2) - (y_1, y_2)\| = |x_1 - y_1| + |x_2 - y_2|$. Consider f in E such that f((x, 1 - x)) = x(1 - x) for all x. f now belongs to D_1 , and $d_1(u, v) \ge 1/4 > 0$ for $u = 1/2 \,\delta_0 + 1/2 \,\delta_1$ and $v = \delta_{1/2}$. One can show that $(\Delta(X), d_1)$ is a compact metric space in this case (see Proposition 2.16 later), and for applications in this paper d_1 will be a particularly useful distance whenever X is a simplex $\Delta(K)$ endowed with $\|x - y\| = \sum_{k \in K} |x^k - y^k|$.

Furthermore it is clear that the Kantorovich-Rubinstein metric on $\Delta(X)$ only depends on the restriction of the norm $\|.\|$ on the set X. Especially if X is finite and $\|x - x'\| = 2$ for all $x \neq x' \in X$, then for all $u, v \in \Delta(X)$ we have $d_{KR}(u, v) = \|u - v\|_1$. This is not the case when considering the metric d_1 . Two norms on V giving the same metric on X may lead to different pseudo-metrics on $\Delta(X)$. We consider in the next example different norms on the Euclidean space \mathbb{R}^K .

Example 2.3. We consider $V = \mathbb{R}^K$, $X = \{e_1, ..., e_K\}$ the set of canonical vectors of V and a norm such that for all $k \neq k'$, $||e_k - e_{k'}|| = 2$. We know that d_1 is smaller than the Kantorovich-Rubinstein metric, so for all $u \in \Delta(X)$ and $v \in \Delta(X)$, we have $d_1(u, v) \leq ||u - v||_1$.

We first consider the particular case of the norm defined by $||x-y|| = 2^{1-\frac{1}{p}} ||x-y||_p$ where $||x-y||_p = \left(\sum_{k=1}^K |x_k - y_k|^p\right)^{1/p}$ is the usual L^p -norm on \mathbb{R}^K , with p a fixed positive integer. Given $u, v \in \Delta(X)$, the function f defined on K by f(k) = 1 if $u(k) \ge v(k)$, and by f(k) = -1 if u(k) < v(k), satisfies $u(f) - v(f) = \sum_{k \in K} |u(k) - v(k)| = ||u - v||_1$. Moreover for all $a \ge 0$, $b \ge 0$ and $k, k' \in K$ such that $k \ne k'$, we have:

$$af(k) - bf(k') \le a + b \le \frac{2}{2^{1/p}} (a^p + b^p)^{1/p} = ||ae_k - be_{k'}||,$$

and $af(k) - bf(k) \leq |a - b| \leq |a - b| \frac{2}{2^{1/p}}$. Therefore f is in D_1 and $d_1(u, v) = ||u - v||_1$, independently of p (one can show that similarly, the same result holds for the case $p = +\infty$, i.e. where $||x - y|| = 2||x - y||_{\infty}$).

Nevertheless the inequality $d_1(u, v) \leq ||u-v||_1$ may be strict as in the following example. We consider the case K = 3 and given a vector $(x_1, x_2, x_3) \in \mathbb{R}^3$, we define the norm $||(x_1, x_2, x_3)|| = \max(|x_1| + |x_2|, 2|x_3|)$, which satisfies $||e_1 - e_2|| =$

¹More generally one can show that if there exists $x \neq 0$ such that the segment [0, x] is included in X, then all f in D_1 are linear on [0, x] and $d_1(\delta_{x/2}, 1/2 \ \delta_0 + 1/2 \ \delta_x) = 0$, so d_1 is not a distance. In the case X = [-1, 1] with the usual norm, one can show that $D_1 = \{f \in E, \exists \alpha, \beta \geq 0 \text{ s.t. } f(x) = \alpha x \text{ for } x \in [-1, 0] \text{ and } f(x) = \beta x \text{ for } x \in [0, 1]\}.$

 $||e_2 - e_3|| = ||e_3 - e_1|| = 2$. Let f be a function in D_1 , we have in particular the following constraints:

$$\begin{aligned} \forall a, b \ge 0 \quad af(e_3) - bf(e_1) &\le \|(-b, 0, a)\| = \max(2a, b) \\ \forall a \ge 0 \quad af(e_2) &\le \|(0, a, 0)\| = a. \end{aligned}$$

Let u = (0, 1/2, 1/2), v = (1, 0, 0) and $f \in D_1$, then

$$u(f) - v(f) = \frac{1}{2}f(e_2) + \frac{1}{2}f(e_3) - f(e_1) \le \frac{1}{2} + \max(2/2, 1) = \frac{3}{2}$$

We deduce that $d_1(u,v) \leq \frac{3}{2} < ||u-v||_1$. In fact one can show that $d_1(u,v) = \frac{3}{2}$ by considering the function defined by $f(e_1) = 0$, $f(e_2) = 1$ and $f(e_3) = 2$. \Box

We now give other expressions for the pseudo-distance d_1 .

Definition 2.4.

a

$$d_2(u, v) = \sup_{(f,g) \in D_2} u(f) + v(g),$$

where $D_2 = \{(f,g) \in E \times E, \forall x, y \in X, \forall a, b \ge 0, af(x) + bg(y) \le ||ax - by||\}.$

Definition 2.5.

$$d_2^+(u,v) = \inf_{\varepsilon > 0} d_2^\varepsilon(u,v), \text{ where } d_2^\varepsilon(u,v) = \sup_{(f,g) \in D_2^\varepsilon} u(f) + v(g)$$

 $and \; \forall \varepsilon > 0, \;\; D_2^\varepsilon = \{(f,g) \in E \times E, \forall x,y \in X, \forall a,b \in [0,1], \; af(x) + bg(y) \le \varepsilon + \|ax - by\|\}.$

Definition 2.6.

$$d_3(u,v) = \inf_{\gamma \in \mathcal{M}_3(u,v)} \int_{X^2 \times [0,1]^2} \|\lambda x - \mu y\| d\gamma(x,y,\lambda,\mu),$$

where $\mathcal{M}_3(u, v)$ is the set of finite positive measures on $X^2 \times [0, 1]^2$ such that for each f in E:

$$\int_{(x,y,\lambda,\mu)\in X^2\times[0,1]^2}\lambda f(x)d\gamma(x,y,\lambda,\mu) = u(f) \ and \ \int_{(x,y,\lambda,\mu)\in X^2\times[0,1]^2}\mu f(y)d\gamma(x,y,\lambda,\mu) = v(f)$$

Theorem 2.7. For all u and v in $\Delta(X)$, $d_1(u, v) = d_2(u, v) = d_2^+(u, v) = d_3(u, v)$.

The proof of Theorem 2.7 is postponed to the next subsection. Recall now the standard Kantorovich duality formula from optimal transport: for any u, v in $\Delta(X)$, let $\Pi(u, v)$ denote the set of transference plans, or couplings, of u and v, that is the set of probability distributions over $X \times X$ with first marginal u and second marginal v. The standard duality formula for d_{KR} reads (see for instance Villani 2003, p.207):

$$d_{KR}(u,v) = \sup_{f \in E_1} |u(f) - v(f)| = \min_{\gamma \in \Pi(u,v)} \int_{(x,y) \in X \times X} ||x - y|| \, d\gamma(x,y)$$

We will use here in subsection 2.3 the duality result $d_1 = d_3$ to prove our main duality formula (Theorem 2.19), in the case where X is a simplex. Compared to the formula for d_{KR} , our proof requires to handle 2 extra variables λ and μ corresponding to a and b in the definition of D_1 (see Remark 2.11). This have led to the introduction of the intermediary quantity d_2^+ . Also, our duality formula Theorem 2.19, as well as our characterization via disintegrations in Theorem 2.22, only applies to probabilities with finite support (see Remark 2.23).

2.2 Proof of Theorem 2.7

The proof is split into several parts.

Proposition 2.8. $d_1 = d_2 = d_2^+$.

It is plain that $d_1 \leq d_2 \leq d_2^+$, so all we have to prove is $d_2^+ \leq d_1$. We start with a lemma.

Lemma 2.9. Fix $\varepsilon > 0$, and let f in E be such that: $\forall x \in X, \forall a \in [0,1], af(x) \le \varepsilon + a ||x||$. Define \hat{f} by:

$$\forall y \in X, \ \hat{f}(y) = \inf_{a \in [0,1], b \in (0,1], x \in X} \frac{1}{b} \left(\varepsilon + \|ax - by\| - af(x) \right).$$

Then for each y in X, $-\|y\| \leq \hat{f}(y) \leq -f(y) + \varepsilon$. Moreover $\hat{f} \in E_1$, and:

$$\forall x \in X, \forall y \in X, \forall a \in [0,1], \forall b \in [0,1], \ a\hat{f}(x) - b\hat{f}(y) \le a\varepsilon + \|by - ax\|.$$

Proof of Lemma 2.9: By assumption on f, we have for all y in X, a in [0, 1], b in (0, 1], x in X: $\frac{1}{b}(\varepsilon + ||ax - by|| - af(x)) \ge \frac{1}{b}(-a||x|| + ||ax - by||) \ge -||y||$. In the definition of $\hat{f}(y)$, considering a = b = 1 and x = y yields $\hat{f}(y) \le -f(y) + \varepsilon$.

Fix x and y in X, a and b in [0, 1]. We have:

$$a\hat{f}(x) - b\hat{f}(y) = a \inf_{a',b',x'} \frac{1}{b'} \left(\varepsilon + \|a'x' - b'x\| - a'f(x')\right) -b \inf_{a'',b'',x''} \frac{1}{b''} \left(\varepsilon + \|a''x'' - b''y\| - a''f(x'')\right).$$

If a = 0, then the inequality $\hat{f}(y) \ge -\|y\|$ leads to $-b\hat{f}(y) \le b\|y\|$. If b = 0, choose a' = 0, b' = 1 and x' = x to get $a\hat{f}(x) \le a\varepsilon + \|ax\|$.

If ab > 0, given $\eta > 0$, choose a'', b'', $x'' \eta$ -optimal in the second infimum. We can define x' = x'', and choose $a' \in [0, 1]$ and $b' \in (0, 1]$ such that $\frac{a'}{b'} = \frac{b}{a} \frac{a''}{b''}$. We obtain:

$$\begin{aligned} a\hat{f}(x) - b\hat{f}(y) &\leq b\eta + (\frac{a}{b'} - \frac{b}{b''})\varepsilon + (\|\frac{a''}{b''}bx'' - ax\| - \|\frac{a''}{b''}bx'' - by\|) \\ &\leq b\eta + (\frac{a}{b'} - \frac{b}{b''})\varepsilon + \|ax - by\|. \end{aligned}$$

If a = b > 0, choose a' = a'' and b' = b'' to obtain: $\hat{f}(x) - \hat{f}(y) \le ||x - y||$ and therefore \hat{f} is 1-Lipschitz.

Otherwise, we distinguish two cases. If $\frac{a}{b}b'' \leq 1$, we define $b' = \frac{a}{b}b''$ and a' = a''and we get $a\hat{f}(x) - b\hat{f}(y) \leq b\eta + ||ax - by||$. If $\frac{a}{b}b'' > 1$, we define b' = 1 and $a' = \frac{a''b}{b''a} \in [0,1]$ and obtain $a\hat{f}(x) - b\hat{f}(y) \leq b\eta + a\varepsilon + ||ax - by||$. Thus for all $\eta > 0$, we have

$$a\hat{f}(x) - b\hat{f}(y) \leq b\eta + a\varepsilon + ||ax - by||,$$

and therefore $a\hat{f}(x) - b\hat{f}(y) \le a\varepsilon + ||ax - by||$.

Proof of Proposition 2.8: Fix u and v in $\Delta(X)$, and consider $\varepsilon > 0$. For each (f,g) in D_2^{ε} , we have $-f + \varepsilon \ge \hat{f} \ge g$ and (f,\hat{f}) in D_2^{ε} . We also have $(\hat{f},f) \in D_2^{\varepsilon}$ so iterating the construction, we get $(\hat{f},\hat{f}) \in D_2^{\varepsilon}$, and $-\hat{f} + \varepsilon \ge \hat{f} \ge f$.

Now, $u(f) + v(g) \le u(\hat{f}) + v(\hat{f}) \le -u(\hat{f}) + \varepsilon + v(\hat{f})$. Hence we have obtained:

$$d_2^{\varepsilon}(u,v) \le \varepsilon + \sup_{f \in C_{\varepsilon(u,v)}} -u(f) + v(f),$$

where $C_{\varepsilon(u,v)}$ is the set of functions f in E_1 satisfying:

$$\forall x \in X, \forall y \in X, \forall a \in [0,1], \forall b \in [0,1], af(x) - bf(y) \le a\varepsilon + \|ax - by\| and f(y) \ge -\|y\|$$

For each positive k, one can choose f_k in E_1 achieving the above supremum for $\varepsilon = 1/k$. Taking a limit point of $(f_k)_k$ yields a function f in D_1 such that: $-u(f) + v(f) \ge d_2^+(u, v)$. The function $f^* = -f$ is in D_1 and satisfies $u(f^*) - v(f^*) \ge d_2^+(u, v)$, and the proof of Proposition 2.8 is complete. \Box

Proposition 2.10. $d_2^+ \geq d_3$.

Proof: The proof is based on (a corollary of) Hahn-Banach theorem. Define: $H = \mathcal{C}(X^2 \times [0, 1]^2)$ and

$$L = \{\varphi \in H, \exists f, g \in \mathcal{C}(X) \ s.t. \ \forall x, y \in X, \forall \lambda, \mu \in [0, 1], \varphi(x, y, \lambda, \mu) = \lambda f(x) + \mu g(y) \}$$

H is endowed with the uniform norm and *L* is a linear subspace of *H*. Note that the unique constant mapping in *L* is 0. Fix *u* and *v* in $\Delta(X)$, and let *r* be the

linear form on L defined by $r(\varphi) = u(f) + v(g)$, where $\varphi(x, y, \lambda, \mu) = \lambda f(x) + \mu g(y)$ for all x, y, λ, μ .

Fix now $\varepsilon > 0$, and put:

$$U_{\varepsilon} = \{\varphi \in H, \forall x, y \in X, \forall \lambda, \mu \in [0, 1], \varphi(x, y, \lambda, \mu) \leq \|\lambda x - \mu y\| + \varepsilon\}.$$

We have:

$$\sup_{\varphi \in L \cap U_{\varepsilon}} r(\varphi) = d_2^{\varepsilon}(u, v).$$

 U_{ε} is a convex subset of H which is radial at 0, in the sense that: $\forall \varphi \in H, \exists \delta > 0$ such that $t\varphi \in U_{\varepsilon}$ as soon as $|t| \leq \delta$. By a corollary of Hahn-Banach theorem (see theorem 6.2.11 p.202 in Dudley, [16]), r can be extended to a linear form on H such that:

$$\sup_{\varphi \in U_{\varepsilon}} r(\varphi) = d_2^{\varepsilon}(u, v).$$

Given $\varphi \in H$, we have $\varepsilon \varphi / \|\varphi\|_{\infty} \in U_{\varepsilon}$, which implies that $r(\varphi) \leq \|\varphi\|_{\infty} d_{2}^{\varepsilon}(u, v) / \varepsilon$, so that r belongs to H'. And if $\varphi \geq 0$, we have $t\varphi \in U_{\varepsilon}$ if $t \leq 0$, so that $r(\varphi) \geq d_{2}^{\varepsilon}(u, v) / t$ for all $t \leq 0$ and $r(\varphi) \geq 0$. By Riesz Theorem, r can be represented by a positive finite measure γ on $X^{2} \times [0, 1]^{2}$.

Given f in E, one can consider $\varphi_f \in L$ defined by $\varphi_f(x, y, \lambda, \mu) = \lambda f(x)$. $r(\varphi_f) = \gamma(\varphi_f)$ gives: $u(f) = \int_{(x,y,\lambda,\mu)\in X^2\times[0,1]^2} \lambda f(x)d\gamma(x,y,\lambda,\mu)$, and similarly $v(f) = \int_{(x,y,\lambda,\mu)\in X^2\times[0,1]^2} \mu f(y)d\gamma(x,y,\lambda,\mu)$, and we obtain that $\gamma \in \mathcal{M}_3(u,v)$.

Since $\gamma \geq 0$, we have $\sup_{\varphi \in U_{\varepsilon}} r(\varphi) = r(\varphi^*)$ where $\varphi^*(x, y, \lambda, \mu) = \|\lambda x - \mu y\| + \varepsilon$. We get $d_2^{\varepsilon}(u, v) = \int_{X^2 \times [0,1]^2} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \varepsilon \gamma(X^2 \times [0,1]^2)$, so

$$d_2^{\varepsilon}(u,v) \ge \int_{X^2 \times [0,1]^2} \|\lambda x - \mu y\| d\gamma(x,y,\lambda,\mu) \ge d_3(u,v).$$

Remark 2.11. The proof given here uses elements of the proof of the standard Kantorovich duality formula in Dudley ([16], see Lemma 11.8.5 p.423). However the arguments need to be more sophisticated here. In particular, there is no need in the standard duality for the extra variables λ and μ , and the analogs of our sets H and L are $\hat{H} = \mathcal{C}(X^2)$ and $\hat{L} = \{\varphi \in \hat{H}, \exists f, g \in \mathcal{C}(X) \text{ s.t. } \forall x, y \in X, \varphi(x, y) =$ $f(x) + g(y)\}$. And it is enough to define $\hat{U} = \{\varphi \in \hat{H}, \forall x, y \in X, \varphi(x, y) < || x - y ||\}$. \hat{U} is convex and open, so radial at any of his element, and $\hat{U} \cap \hat{L}$ is not empty. In the present setup, if we simply define $U = \{\varphi \in H, \forall x, y \in X, \forall \lambda, \mu \in$ $[0,1], \varphi(x, y, \lambda, \mu) < ||\lambda x - \mu y||\}$, we have $U \cap L = \emptyset$, hence a problem. These considerations have led to the introduction of the sets U_{ε} and the intermediate distance d_2^+ beforehand. \Box

It is now easy to conclude the proof of Theorem 2.7.

Lemma 2.12. $d_3 \ge d_2$.

Proof: Fix $(f,g) \in D_2$ and $\gamma \in \mathcal{M}_3(u,v)$.

$$\begin{split} u(f) + v(g) &= \int_{X^2 \times [0,1]^2} \lambda f(x) d\gamma(x, y, \lambda, \mu) + \int_{X^2 \times [0,1]^2} \mu g(y) d\gamma(x, y, \lambda, \mu) \\ &= \int_{X^2 \times [0,1]^2} (\lambda f(x) + \mu g(y)) d\gamma(x, y, \lambda, \mu) \\ &\leq \int_{X^2 \times [0,1]^2} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu). \end{split}$$

2.3 The case of probabilities over a simplex

The case where X itself is a probability space is interesting, and we assume here that $X = \Delta(K)$ is a simplex, where K is a non empty finite set. We use $\|p\| = \sum_{k} |p^{k}|$ for every vector $p = (p^{k})_{k \in K}$ in \mathbb{R}^{K} , and view X as the set of vectors in \mathbb{R}^{K}_{+} with norm 1.

$$X = \{ p = (p^k)_{k \in K} \in I\!\!R_+^K, \sum_{k \in K} p^k = 1 \}.$$

Recall that for u and v in $\Delta(X)$, we have $d_1(u, v) = \sup_{f \in D_1} |u(f) - v(f)|$, where $D_1 = \{f \in E, \forall x, y \in X, \forall a, b \ge 0, af(x) - bf(y) \le ||ax - by||\}.$

We now introduce an alternative definition of d_1 using "non revealing game functions". These functions come from the theory of repeated games with incomplete information à la Aumann Maschler [5], and the interest for the distance d_0 emerged several years ago while doing research on Markov decision processes with partial observation and repeated games with an informed controller (see Renault [32] and [33]).

Given a collection of matrices $(G^k)_{k \in K}$ (all of the same finite size $I \times J$) indexed by K and with values in [-1, 1], we define the "non revealing function" f in $\mathcal{C}(X)$ by:

$$\begin{aligned} \forall p \in X, f(p) &= \operatorname{Val}\left(\sum_{k \in K} p^k G^k\right), \\ &= \max_{x \in \Delta(I)} \min_{y \in \Delta(J)} \sum_{i \in I, j \in J} x(i) y(j) \left(\sum_{k \in K} p^k G^k(i, j)\right), \\ &= \min_{y \in \Delta(J)} \max_{x \in \Delta(I)} \sum_{i \in I, j \in J} x(i) y(j) \left(\sum_{k \in K} p^k G^k(i, j)\right). \end{aligned}$$

Here Val denotes the minmax value of a matrix, and f(p) is the minmax value of the average matrix $\sum_{k} p^{k} G^{k}$. The set of all such non revealing functions f, where I, J and $(G^{k})_{k \in K}$ vary, is denoted by D_{0} . Clearly, all affine functions from X to [-1, 1] belong to D_0 . Using the Stone-Weierstrass theorem, one can show that the set of *all* non revealing functions is dense in $\mathcal{C}(X)$ (see proposition 5.1. chapter VI p.357 in Mertens *et al.* [29] and Lemma 2.13 below). However, we only consider here non revealing functions defined by matrices with values in [-1, 1], and D_0 is *not* dense in the set of continuous functions from X to [-1, 1]. As an example, consider the case where $K = \{1, 2\}$ and f in E is piecewise-linear with f(1, 0) = f(0, 1) = 0 and f(1/2, 1/2) = 1. Notice that if a function g in D_0 is such that g(1/2, 1/2) = 1, then necessarily the values of the two matrix games G^1 and G^2 are also equal to 1 since it is the maximum value for all possible payoffs, and necessarily g(1, 0) = g(0, 1) = 1. Therefore f is not in D_0 . In fact f is 1-Lipschitz, however 2f(1/2, 1/2) - f(1, 0) = 2 > ||2(1/2, 1/2) - (1, 0)|| = 1, so it is not in D_1 which we will see later contains D_0 (see Lemma 2.14).

Lemma 2.13. If f, g belong to D_0 and $\lambda \in [0, 1]$, then -f, $\sup\{f, g\}$, $\inf\{f, g\}$ and $\lambda f + (1 - \lambda)g$ are in D_0 . The linear span of D_0 is dense in $\mathcal{C}(X)$.

Proof: The proof can be easily deduced from proposition 5.1. page 357 in Mertens *et al.* [29], chapter VI part B. For instance, let f and g in D_0 be respectively defined by the collections of matrices $(G^k)_{k \in K}$ with size $I_1 \times J_1$ and $(H^k)_{k \in K}$ with size $I_2 \times J_2$.

Defining for each k, i_1, j_1 : $G'^k(i_1, j_1) = -G^k(j_1, i_1)$ yields a family of matrices $(G'^k)_k$ with size $J_1 \times I_1$ inducing -f. So $-f \in D_0$.

To get that $\sup\{f, g\}$ belongs to D_0 , one can assume w.l.o.g. that $I_1 \cap I_2 = J_1 \cap J_2 = \emptyset$. Set $I = I_1 \cup I_2$ and $J = J_1 \times J_2$. Define for each k the matrix game L^k in $\mathbb{R}^{I \times J}$ by $L^k(i, (j_1, j_2)) = G^k(i, j_1)$ if $i \in I_1$, $L^k(i, (j_1, j_2)) = H^k(i, j_2)$ if $i \in I_2$. Then for each p in X, we have $\operatorname{Val}(\sum_k p^k L^k) = \sup\{f(p), g(p)\}$, so that $\sup\{f, g\} \in D_0$.

Lemma 2.14. The closure of D_0 is D_1 .

Proof: We first show that $D_0 \subset D_1$. Let I and J be finite sets, and $(G^k)_{k \in K}$ be a collection of $I \times J$ -matrices with values in [-1, 1]. Consider p and q in X and a and b non negative. Then for all i and j:

$$\left|\sum_{k} p^{k} a G^{k}(i,j) - \sum_{k} q^{k} b G^{k}(i,j)\right| \leq \sum_{k} |ap^{k} - bq^{k}| = ||ap - bq||.$$

As a consequence,

$$a \operatorname{Val}\left(\sum_{k \in K} p^{k} G^{k}\right) - b \operatorname{Val}\left(\sum_{k \in K} q^{k} G^{k}\right) = \operatorname{Val}\left(\sum_{k \in K} a p^{k} G^{k}\right) - \operatorname{Val}\left(\sum_{k \in K} b q^{k} G^{k}\right)$$
$$\leq \|ap - bq\|$$

We now show that the closure of D_0 contains D_1 . Consider f in D_1 , in particular we have $||f||_{\infty} \leq 1$. Let p and q be distinct elements in X, we define Y

as the linear span of p and q, and φ from Y to \mathbb{R} by : $\varphi(\lambda p + \mu q) = \lambda f(p) + \mu f(q)$ for all reals λ and μ .

If $\lambda \geq 0$ and $\mu \geq 0$, we have $\varphi(\lambda p + \mu q) \leq \lambda + \mu = ||\lambda p + \mu q||$. If $\lambda \geq 0$ and $\mu \leq 0$, we use the definition of D_1 to get: $\varphi(\lambda p + \mu q) \leq ||\lambda p + \mu q||$. As a consequence, φ is a linear form with norm at most 1 on Y. By Hahn-Banach theorem, it can be extended to a linear mapping on \mathbb{R}^K with the same norm, and we denote by g the restriction of this mapping to X. g is affine with $g(p) = \varphi(p) = f(p)$ and $g(q) = \varphi(q) = f(q)$. Moreover, for each r in X, we have $|g(r)| \leq ||r|| = 1$. As a consequence g belongs to D_0 .

Because D_0 is stable under the sup and inf operations, we can use the Stone-Weierstrass theorem (see for instance lemma A7.2 in Ash [2] p.392) to conclude that f belongs to the closure of D_0 .

Definition 2.15. Given u and v in $\Delta(X)$, define:

$$d_0(u, v) = \sup_{f \in D_0} u(f) - v(f).$$

Proposition 2.16. d_0 is a distance on $\Delta(X)$ metrizing the weak-* topology. Moreover $d_0 = d_1 = d_2 = d_3$.

Proof: $d_0 = d_1 = d_2 = d_3$ follows from Lemma 2.14 and Theorem 2.7. Because the linear span of D_0 is dense in $\mathcal{C}(X)$, we obtain the separation property and d_0 is a distance on $\Delta(X)$. Because $D_0 \subset D_1 \subset E_1$, we have $d_0 = d_1 \leq d_{KR}$. Since $(\Delta(X), d_{KR})$ is a compact metric space, the identity map $(\Delta(X), d_{KR})$ to $(\Delta(X), d_0)$ is bicontinuous, and we obtain that $(\Delta(X), d_0)$ is a compact metric space and d_0 and d_{KR} are equivalent. (see for instance proposition 2 page 138 [4].

Remarks 2.17.

1) Let LF be the set of linear forms on $(\mathbb{R}^K, \|.\|_1)$ with norm at most 1. In Lemma 2.14, D_0 can be replaced by the lattice generated by the restrictions to $\Delta(K)$ of the elements of LF.

2) One can show that allowing for infinite sets I, J in the definition of D_0 (still assuming that all games $\sum_k p^k G^k$ have a value) would not change the value of d_0 .

From now on, we just write $d_*(u, v)$ for the distance $d_0 = d_1 = d_2 = d_3$ on $\Delta(X)$. Elements of X can be viewed as elements of $\Delta(X)$ (using Dirac measures), and for p, q in X, we have: $d_{KR}(\delta_p, \delta_q) = d_*(\delta_p, \delta_q) = ||p - q||$. We now present a dual formulation for our distance, in the spirit of Kantorovich duality formula from optimal transport. We will concentrate on probabilities on X with finite support, and denote by $Z = \Delta_f(X)$ the set of such probabilities.

Definition 2.18. Let u and v be in Z with respective supports U and V. We define the set $\mathcal{M}_4(u, v) =$

$$\left\{ (\alpha,\beta) \in (I\!\!R_+^{U\times V})^2, s.t. \ \forall x \in U, \sum_{y' \in V} \alpha(x,y') = u(x) \text{ and } \forall y \in V, \sum_{x' \in U} \beta(x',y) = v(y) \right\}.$$

And we put $d_4(u,v) = \inf_{(\alpha,\beta) \in \mathcal{M}_4(u,v)} \sum_{(x,y) \in U \times V} \|x\alpha(x,y) - y\beta(x,y)\|.$

 $\mathcal{M}_4(u, v)$ is the set of couples (α, β) of probability measures on $U \times V$ such that the first marginal of α is u and the second marginal of β is v. Notice that diagonal elements in $\mathcal{M}_4(u, v)$ coincide with elements of $\Pi(u, v)$, i.e. with probability distributions over $X \times X$ with first marginal u and second marginal v. The set $\mathcal{M}_4(u, v)$ is simply a polytope in the Euclidean space $(I\!\!R^{U \times V})^2$, so the infimum in the definition of $d_4(u, v)$ is achieved. Recall that $D_1 = \{f \in E, \forall x, y \in$ $X, \forall a, b \geq 0, af(x) - bf(y) \leq ||ax - by||\}$. The next theorem is the main result of this section.

Theorem 2.19. (Duality formula) Let u and v be in Z with respective supports U and V.

$$d_*(u,v) = \sup_{f \in D_1} |u(f) - v(f)| = \min_{(\alpha,\beta) \in \mathcal{M}_4(u,v)} \sum_{(x,y) \in U \times V} ||x\alpha(x,y) - y\beta(x,y)||.$$

The proof is postponed to the next subsection. We conclude this part by a simple but universal property of the distance d_* .

Definition 2.20. Given a finite set S, we define the disintegration, or posterior mapping, ψ_S from $\Delta(K \times S)$ to $\Delta(X)$ by:

$$\psi_S(\pi) = \sum_{s \in S} \pi(s) \delta_{p(s)}$$

where for each s, $\pi(s) = \sum_k \pi(k, s)$ and $p(s) = (p^k(s))_{k \in K} \in X$ is the posterior on K given s (defined arbitrarily if $\pi(s) = 0$): for each k in K, $p^k(s) = \frac{\pi(k,s)}{\pi(s)}$.

 $\psi_S(\pi)$ is a probability with finite support over X. Intuitively, think of a joint variable (k, s) being selected according to π , and an agent just observes s. His knowledge on K is then represented by p(s). And $\psi_S(\pi)$ represents the ex-ante information that the agent will know about the variable k.

 $\Delta(K \times S)$ is endowed as usual with the $\|.\|_1$ norm. One can show that ψ_S is continuous whenever X is endowed with the weak-* topology. Intuitively, $\psi_S(\pi)$ has less information than π , because the agent does not care about s itself but

just on the information about k given by s. So one may hope that the mapping ψ_S is 1-Lipschitz (non expansive) for a well chosen distance on $\Delta(X)$. It is not a *priori* obvious to see whether the Kantorovich-Rubinstein distance d_{KR} has this property, the following example shows that it is not the case.

Example 2.21. Consider the case where $K = \{a, b, c\}$ and $S = \{\alpha, \beta\}$. We denote by π and π' the following laws on $\Delta(K \times S)$:

$$K \begin{pmatrix} S & S \\ \frac{1}{4} & 0 \\ 0 & \frac{1}{2} \\ \frac{1}{4} & 0 \end{pmatrix} \text{ and } \begin{pmatrix} \frac{1}{4} & 0 \\ 0 & \frac{1}{2} \\ 0 & \frac{1}{4} \end{pmatrix}.$$
$$\pi & \pi'$$

Their disintegrations are respectively $\psi_S(\pi) = \frac{1}{2} (1/2, 0, 1/2) + \frac{1}{2} (0, 1, 0)$ and $\psi_S(\pi') = \frac{1}{4}(1, 0, 0) + \frac{3}{4}(0, 2/3, 1/3)$. We define the test function $f : \Delta(K) \to [-1, 1]$ by:

$$f(0,1,0) = \frac{1}{3}, f(1/2,0,1/2) = -\frac{1}{3}, f(1,0,0) = \frac{2}{3}, f(0,2/3,1/3) = 1.$$

f can be extended to a 1-Lispchitz function on the simplex $\Delta(K)$ by the Mc Shane-Whitney extension theorem (Mc Shane, [25]), or directly by $f(p_a, p_b, p_c) = \max\{1 - p_a - |p_b - 2/3| - |p_c - 1/3|, -1/3 + |p_a - p_c|\}$). We have $||\pi - \pi'|| = \frac{1}{2}$ and $d_{KR}(\psi_S(\pi), \psi_S(\pi')) \ge \psi_S(\pi')(f) - \psi_S(\pi)(f) = \frac{11}{12} - 0 > \frac{1}{2}$. So the posterior mapping ψ_S is not 1-Lipschitz from $(\Delta(K \times S), ||.||_1)$ to $(\Delta(X), d_{KR})$.

The next theorem shows that our distance d_* has the fundamental property to make all disintegrations ψ_S non expansive, and is the largest distance to do so.

Theorem 2.22. For each finite set S, the mapping ψ_S is 1-Lipschitz from $(\Delta(K \times S), \|.\|_1)$ to $(\Delta_f(X), d_*)$. Moreover, d_* is the largest distance on Z having this property:

$$\forall u, u' \in Z, \ d_*(u, u') = \min\{\|\pi - \pi'\|_1, \ s.t. \ \exists S \ finite, \ \psi_S(\pi) = u, \ \psi_S(\pi') = u'\}.$$

The proof is simple after having obtained the duality formula. **Proof:** First fix S and π , π' in $\Delta(K \times S)$. Write $u = \psi_S(\pi)$, $u' = \psi_S(\pi')$. For any f in D_1 , we have:

$$u(f) - u'(f) = \sum_{s \in S} (\pi(s)f(p(s)) - \pi'(s)f(p'(s)))$$

$$\leq \sum_{s \in S} \|\pi(s)p(s) - \pi'(s)p'(s)\|$$

$$\leq \sum_{s \in S} \|(\pi(k,s))_k - (\pi'(k,s))_k\|$$

$$\leq \sum_{s \in S} \sum_{k \in K} |\pi(k,s) - \pi'(k,s)| = \|\pi - \pi'\|_1.$$

So $d_*(u, u') \leq ||\pi - \pi'||_1$, and ψ_S is 1-Lipschitz.

Let now u and v be in Z with respective supports U and V. Using the duality formula of Theorem 2.19, one can find $(\alpha, \beta) \in \mathcal{M}_4(u, v)$ such that

$$d_*(u,v) = \sum_{(x,y)\in U\times V} \|\alpha(x,y)x - \beta(x,y)y\|$$

Define $S = U \times V$ and $\pi, \pi' \in \Delta(K \times S)$ by $\pi(k, (x, y)) = x(k)\alpha(x, y)$ and $\pi'(k, (x, y)) = y(k)\beta(x, y)$. By definition of $\mathcal{M}_4(u, v), \pi$ and π' are probabilities and

$$\begin{aligned} \|\pi - \pi'\|_{1,K\times S} &= \sum_{k\in K, (x,y)\in U\times V} |x(k)\alpha(x,y) - y(k)\beta(x,y)| \\ &= \sum_{(x,y)\in U\times V} \|\alpha(x,y)x - \beta(x,y)y\|.\Box \end{aligned}$$

Finally notice that considering infinite sets S dramatically changes the picture, as shown by the following simple example (communicated by Filippo Santambrogio).

Remark 2.23. Fix $K = \{a, b\}$ and S = [0, 1], and define for each π in $\Delta(K \times S)$, the image $\psi_S(\pi)$ in $\Delta(X)$ by: $\psi_S(\pi)(f) = \int_{k,s} f(p(s))d\pi(k,s)$ for all f in $\mathcal{C}(X)$ (here again, $X = \Delta(K)$ and p(s) is the posterior on K given S). $\Delta(K \times S)$ and $\Delta(X)$ are endowed with weak-* topologies.

Consider the uniform probability π over $K \times S$, then $\psi_S(\pi) = \delta_{\frac{1}{2}a + \frac{1}{2}b}$ is the Dirac measure on $\frac{1}{2}a + \frac{1}{2}b$. We approximate π by considering finer and finer grids of the unit interval. For each positive integer n, partition [0,1) into $A_n = \bigcup_{k=0}^{n-1} [\frac{2k}{2n}, \frac{2k+1}{2n}]$ and $B_n = \bigcup_{k=0}^{n-1} [\frac{2k+1}{2n}, \frac{2k+2}{2n}]$. And we can define π_n by first choosing s in S according to the Lebesgue measure, then set k = a if $s \in A_n$ and set k = b if $s \in B_n$. Knowing s perfectly determines k here, and $\psi_S(\pi_n) = \frac{1}{2}\delta_a + \frac{1}{2}\delta_b$. However π_n converges to π , so ψ_S is not even continuous.

2.4 Proof of the duality formula

Let u and v be in $\Delta(X)$, and denote by U and V the respective supports of uand v. We write $S = X^2 \times [0, 1]^2$, and we start with a lemma, where no finiteness assumption on U or V is needed. Recall that $\mathcal{M}_3(u, v)$ is the set of finite positive measures on S satisfying for each f in E:

$$\int_{(x,y,\lambda,\mu)\in S} \lambda f(x) d\gamma(x,y,\lambda,\mu) = u(f), \text{ and } \int_{(x,y,\lambda,\mu)\in S} \mu f(y) d\gamma(x,y,\lambda,\mu) = v(f).$$

Lemma 2.24. For each $\gamma \in \mathcal{M}_3(u, v)$, we have:

$$\int_{S} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) = 2 + \int_{U \times V \times [0,1]^2} \left(\|\lambda x - \mu y\| - \lambda - \mu \right) d\gamma(x, y, \lambda, \mu).$$

Proof: Write $A(\gamma) = \int_S ||\lambda x - \mu y|| d\gamma(x, y, \lambda, \mu)$. Using the definition of $\mathcal{M}_3(u, v)$, we can obtain that $1 = \int_S \lambda \mathbf{1}_{x \in U} d\gamma = \int_S \mu \mathbf{1}_{y \in V} d\gamma$. This implies: $\int_S \lambda \mathbf{1}_{x \notin U} d\gamma = \int_S \mu \mathbf{1}_{y \notin V} d\gamma = 0$, so that $\lambda \mathbf{1}_{x \notin U} = \mu \mathbf{1}_{y \notin V} = 0$ γ . a.s. We can write:

$$\begin{split} A(\gamma) &= \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) \\ &+ \int_{S} \mathbf{1}_{x \notin U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \notin U, y \notin V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) \\ &= \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \notin U, y \in V} \mu d\gamma(x, y, \lambda, \mu) + 0 \\ & \text{We now use } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \lambda d\gamma + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + \int_{S} \mathbf{1}_{x \in U, y \notin V} \lambda d\gamma \text{ and } \mathbf{1} = \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda y\| d\gamma(x, y, \lambda, \mu) + 0 \\ & \int_{S}$$

$$A(\gamma) = 2 + \int_{S} \mathbf{1}_{x \in U, y \in V} \|\lambda x - \mu y\| d\gamma(x, y, \lambda, \mu) - \int_{S} \mathbf{1}_{x \in U, y \in V} \lambda d\gamma - \int_{S} \mathbf{1}_{x \in U, y \in V} \mu d\gamma.$$

We assume in the sequel that U and V are finite, and define $d_5(u, v)$ as follows: Definition 2.25. Define

$$\mathcal{M}_{5}(u,v) = \left\{ (\alpha,\beta) = (\alpha(x,y),\beta(x,y))_{(x,y)\in U\times V} \in \mathbb{R}_{+}^{U\times V} \times \mathbb{R}_{+}^{U\times V}, s.t. \\ \forall x \in U, \sum_{y'\in V} \alpha(x,y') \leq u(x) \text{ and } \forall y \in V, \sum_{x'\in U} \beta(x',y) \leq v(y) \right\}.$$

And $d_{5}(u,v) = \inf_{(\alpha,\beta)\in\mathcal{M}_{5}(u,v)} \left(2 + \sum_{(x,y)\in U\times V} \left(\|x\alpha(x,y) - y\beta(x,y)\| - \alpha(x,y) - \beta(x,y) \right) \right).$

 $\mathcal{M}_5(u,v)$ is a polytope in the Euclidean space $(\mathbb{R}^{U \times V})^2$, so the infimum in the definition of $d_5(u,v)$ is achieved.

Lemma 2.26. $d_3(u, v) \ge d_5(u, v)$.

Proof: Let γ be in $\mathcal{M}_3(u, v)$. Fix for a while (x, y) in $U \times V$, and assume that $\gamma(x, y) > 0$. We define $\gamma(.|x, y)$ the conditional probability on $[0, 1]^2$ given (x, y) by: for all $\varphi \in C([0, 1]^2)$,

$$\int_{[0,1]^2} \varphi(\lambda,\mu) d\gamma(\lambda,\mu|x,y) = \frac{1}{\gamma(x,y)} \int_{(x',y',\lambda,\mu)\in S} \mathbf{1}_{x'=x,y'=y} \varphi(\lambda,\mu) d\gamma(x',y',\lambda,\mu).$$

So that

$$\gamma(x,y)\int_{[0,1]^2} (\|\lambda x - \mu y\| - \lambda - \mu)d\gamma(\lambda,\mu|x,y) = \int_{(\lambda,\mu)\in[0,1]^2} (\|\lambda x - \mu y\| - \lambda - \mu)d\gamma(x,y,\lambda,\mu)d\gamma(x,\mu)d\gamma($$

Define $P(x, y) = \int_{(\lambda, \mu) \in [0,1]^2} \lambda d\gamma(\lambda, \mu | x, y)$ and $Q(x, y) = \int_{(\lambda, \mu) \in [0,1]^2} \mu d\gamma(\lambda, \mu | x, y)$. The mapping $\Psi : (\lambda, \mu) \mapsto \|\lambda x - \mu y\| - \lambda - \mu$ is convex so by Jensen's inequality we get:

$$\int_{(\lambda,\mu)\in[0,1]^2} (\|\lambda x - \mu y\| - \lambda - \mu) d\gamma(\lambda,\mu|x,y) \ge \|xP(x,y) - yQ(x,y)\| - P(x,y) - Q(x,y).$$

Now, by Lemma 2.24,

$$\begin{aligned} A(\gamma) &= 2 + \sum_{x \in U, y \in V} \int_{(\lambda, \mu) \in [0, 1]^2} \left(\|\lambda x - \mu y\| - \lambda - \mu \right) d\gamma(x, y, \lambda, \mu) \\ &= 2 + \sum_{x \in U, y \in V, \gamma(x, y) > 0} \int_{(\lambda, \mu) \in [0, 1]^2} \left(\|\lambda x - \mu y\| - \lambda - \mu \right) d\gamma(x, y, \lambda, \mu) \\ &\geq 2 + \sum_{x \in U, y \in V, \gamma(x, y) > 0} \gamma(x, y) \left(\|x P(x, y) - y Q(x, y)\| - P(x, y) - Q(x, y) \right). \end{aligned}$$

For (x, y) in $U \times V$, define $\alpha(x, y) = \gamma(x, y)P(x, y) \ge 0$ and $\beta(x, y) = \gamma(x, y)Q(x, y) \ge 0$ (with $\alpha(x, y) = \beta(x, y) = 0$ if $\gamma(x, y) = 0$). We get:

$$A(\gamma) \ge 2 + \sum_{x \in U, y \in V} \left(\|x\alpha(x, y) - y\beta(x, y)\| - \alpha(x, y) - \beta(x, y) \right).$$

And we have, for each x in U:

$$\sum_{y \in V} \alpha(x, y) = \sum_{y \in V, \gamma(x, y) > 0} \int_{(\lambda, \mu) \in [0, 1]^2} \lambda d\gamma(x, y, \lambda, \mu)$$
$$\leq \int_{(y, \lambda, \mu) \in X \times [0, 1]^2} \lambda d\gamma(x, y, \lambda, \mu) = u(x).$$

where the last equality comes from the definition of $\mathcal{M}_3(u, v)$. Similarly, for each y in V we can show that $\sum_{x \in U} \beta(x, y) \leq v(y)$, and Lemma 2.26 is proved.

Lemma 2.27. $d_5(u, v) \ge d_4(u, v)$.

Proof: Consider (α^*, β^*) achieving the minimum in the definition of $d_5(u, v)$. Assume that there exists x^* such that $\sum_{y \in V} \alpha(x^*, y) < u(x^*)$. For any x in X and z in \mathbb{R}^K_+ , it is easy to see that the mapping $l : (\alpha \mapsto ||x\alpha - z|| - \alpha)$ is nonincreasing from \mathbb{R}_+ to \mathbb{R} (as the sum of the mappings $l_k : (\alpha \mapsto |\alpha x^k - z^k| - \alpha x^k)$, each l^k being non increasing in α). As a consequence, one can choose any y^* in V and increase $\alpha(x^*, y^*)$ in order to saturate the constraint without increasing the objective. So we can assume without loss of generality that $\sum_{y \in V} \alpha(x^*, y) = u(x^*)$ for all x^* and similarly $\sum_{x \in U} \beta(x, y^*) = v(y^*)$ for all y^* .

Consequently,

$$d_{5}(u,v) = 2 + \sum_{(x,y)\in U\times V} (\|x\alpha^{*}(x,y) - y\beta^{*}(x,y)\| - \alpha^{*}(x,y) - \beta^{*}(x,y))$$

$$= \sum_{(x,y)\in U\times V} \|x\alpha^{*}(x,y) - y\beta^{*}(x,y)\| \ge d_{4}(u,v).$$

Lemma 2.28. $d_4(u, v) \ge d_2(u, v)$.

Proof: Fix $(f,g) \in D_2$ and $(\alpha,\beta) \in \mathcal{M}_4(u,v)$.

$$u(f) + v(g) = \sum_{x \in U} f(x)u(x) + \sum_{y \in Y} g(y)v(y)$$

$$= \sum_{(x,y) \in U \times V} f(x)\alpha(x,y) + g(y)\beta(x,y)$$

$$\leq \sum_{(x,y) \in U \times V} \|\alpha(x,y)x - \beta(x,y)y\| \le d_4(u,v).$$

We have shown that $d_3(u,v) \ge d_5(u,v) \ge d_4(u,v) \ge d_2(u,v) = d_3(u,v) = d_1(u,v)$. This ends the proof of Theorem 2.19.

3 Long-term values for compact non expansive Markov Decision Processes

In this section we consider Markov Decision Processes, or Controlled Markov Chains, with bounded payoffs and transitions with finite support. We will consider two closely related models of MDP and prove in each case the existence and a characterization for a general notion of long-term value.

1) The first model deals with MDP without any explicit action set (hence, payoffs only depend on the current state), such MDP will be called gambling houses using the terminology of gambling theory (see Maitra and Sudderth [24]). We will assume in this setup that the set of states X is metric compact and that the transitions are non expansive with respect to the distance d_{KR} on $\Delta(X)$. Since we only use the distance d_{KR} here, the theorem for the first model, namely Theorem 3.10, does not use the distance for belief spaces studied in section 2.

2) The second model is the standard model of Markov Decision Processes with states, actions, transitions and payoffs, and we will assume that the state space X is a compact subset of a simplex $\Delta(K)$. We will need for this second case an assumption of non expansiveness for the transitions which is closely related to the distance d_* introduced in section 2, see Theorem 3.20 later. The applications in sections 4.1 and 4.2 will be based on this second model.

3.1 Long-term values for Gambling Houses

In this section we consider Markov Decision Processes of the following form. There is a non empty set of states X, a transition given by a multi-valued mapping $F: X \Rightarrow \Delta_f(X)$ with non empty values, and a payoff (or reward) function $r: X \to [0, 1]$. The interpretation is that given an initial state x_0 in X, a decision-maker (or player) has to choose a probability with finite support u_1 in $F(x_0)$, then x_1 is selected according to u_1 and there is a payoff $r(x_1)$. Then the player has to choose u_2 in $F(x_1)$, x_2 is selected according to u_1 and the player receives the payoff $r(x_2)$, etc... Note that there is no explicit action set here, and that the transitions take values in $\Delta_f(X)$ and hence have finite support.

We say that $\Gamma = (X, F, r)$ is a Gambling House. We assimilate the elements in X with their Dirac measures in $\Delta(X)$, and in case the values of F only consist of Dirac measures on X, we view F as a correspondence from X to X and say that Γ is a *deterministic* Gambling House (or a Dynamic Programming problem). In general we write $Z = \Delta_f(X)$, and an element in Z is written $u = \sum_{x \in X} u(x) \delta_x$. The set of stages is $\mathbb{N}^* = \{1, ..., t, ...\}$, and a probability distribution over stages is called an evaluation. Given an evaluation $\theta = (\theta_t)_{t\geq 1}$ and an initial stage x_0 in X, the θ -problem $\Gamma_{\theta}(x_0)$ is the optimization problem defined by a decision-maker starting from x_0 and maximizing the expectation of $\sum_{t\geq 1} \theta_t r(x_t)$.

Formally, we first linearly extend r and F to $\Delta_f(\overline{X})$ by defining for each $u = \sum_{x \in X} u(x)\delta_x$ in Z, the payoff $r(u) = \sum_{x \in X} r(x)u(x)$ and the transition $F(u) = \{\sum_{x \in X} u(x)f(x), s.t. \ f: X \to Z \text{ and } f(x) \in F(x) \ \forall x \in X\}.$

Definition 3.1. The mixed extension of F is the correspondence from Z to itself which associates to every $u = \sum_{x \in X} u(x)\delta_x$ in $\Delta_f(X)$ the image:

$$\hat{F}(u) = \left\{ \sum_{x \in X} u(x) f(x), \text{ s.t. } f: X \to Z \text{ and } f(x) \in \operatorname{conv} F(x) \ \forall x \in X \right\}.$$

The graph of \hat{F} is the convex hull of the graph of F. Moreover \hat{F} is an affine correspondence, as shown by the lemma below.

Lemma 3.2. $\forall u, u' \in Z, \forall \alpha \in [0, 1], \hat{F}(\alpha u + (1 - \alpha)u') = \alpha \hat{F}(u) + (1 - \alpha)\hat{F}(u').$

Proof: The \subset part is clear. To see the reverse inclusion, let $v = \alpha \sum_{x \in X} u(x) f(x) + (1-\alpha) \sum_{x \in X} u'(x) f'(x)$ be in $\alpha \hat{F}(u) + (1-\alpha) \hat{F}(u')$, with transparent notations. Define

$$g(x) = \frac{\alpha u(x)f(x) + (1 - \alpha)u'(x)f'(x)}{\alpha u(x) + (1 - \alpha)u'(x)},$$

for each x such that the denominator is positive. Then $g(x) \in \operatorname{conv} F(x)$, and

$$v = \sum_{x \in X} (\alpha u(x) + (1 - \alpha)u'(x)) g(x) \in \hat{F}(\alpha u + (1 - \alpha)u').$$

Definition 3.3. A pure play, or deterministic play, at x_0 is a sequence $\sigma = (u_1, ..., u_t, ...) \in Z^{\infty}$ such that $u_1 \in F(x_0)$ and $u_{t+1} \in F(u_t)$ for each $t \geq 1$. A play, or mixed play, at x_0 is a sequence $\sigma = (u_1, ..., u_t, ...) \in Z^{\infty}$ such that $u_1 \in \operatorname{conv} F(x_0)$ and $u_{t+1} \in \hat{F}(u_t)$ for each $t \geq 1$. We denote by $\Sigma(x_0)$ the set of mixed plays at x_0 . A pure play is a particular case of a mixed play. Mixed plays corresponds to situations where the decision-maker can select, at every stage t and state x_{t-1} , randomly the law u_t of the new state. A mixed play at x_0 naturally induces a probability distribution over the set $(X \times \Delta_f(X))^{\infty}$ of sequences $(x_0, u_0, x_1, u_1, ...)$, where X and Z are endowed with the discrete σ -algebra and $(X \times \Delta_f(X))^{\infty}$ is endowed with the product σ -algebra.

Definition 3.4. Given an evaluation θ , the θ -payoff of a play $\sigma = (u_1, ..., u_t, ...)$ is defined as: $\gamma_{\theta}(\sigma) = \sum_{t \geq 1} \theta_t r(u_t)$, and the θ -value at x_0 is:

$$v_{\theta}(x_0) = \sup_{\sigma \in \Sigma(x_0)} \gamma_{\theta}(\sigma).$$

It is easy to see that the supremum in the definition of v_{θ} can be taken over the set of pure plays at x_0 . We have the following recursive formula. For each evaluation $\theta = (\theta_t)_{t \ge 1}$ such that $\theta_1 < 1$, we denote by θ^+ the "shifted" evaluation $\left(\frac{\theta_{t+1}}{1-\theta_1}\right)_{t \ge 1}$. We extend linearly v_{θ} to Z, so that the recursive formula reads:

$$\forall \theta \in \Delta(\mathbb{N}^*), \forall x \in X, \ v_{\theta}(x) = \sup_{u \in \operatorname{conv} F(x)} \left(\theta_1 r(u) + (1 - \theta_1) v_{\theta^+}(u) \right).$$

And by linearity the supremum can be taken over F(x). It is also easy to see that for all evaluations θ and initial states x, we have the inequality:

$$|v_{\theta}(x) - \sup_{u \in F(x)} v_{\theta}(u)| \le \theta_1 + \sum_{t \ge 2} |\theta_t - \theta_{t-1}|.$$

$$\tag{1}$$

In this paper, we are interested in the limit behavior when the decision-maker is very patient. Given an evaluation θ , we define the total variation of θ by:

$$TV(\theta) = \sum_{t \ge 1} |\theta_{t+1} - \theta_t|.$$

The decision-maker is considered as patient whenever $TV(\theta)$ is small, so $TV(\theta)$ may be seen as the impatience of θ (see Renault [34] and Sorin [41] p. 105). When $\theta = (\theta_t)_{t\geq 1}$ is non increasing, then $TV(\theta)$ is just θ_1 . A classic example is when $\theta = \frac{1}{n} \sum_{t=1}^{n} \delta_t$, the value v_{θ} is just denoted v_n and the evaluation corresponds to the average payoff from stage 1 to stage n. In this case $TV(\theta) = 1/n \xrightarrow[n\to\infty]{} 0$. We also have $TV(\theta) = 1/n$ if $\theta = \sum_{t=m}^{m+n} \frac{1}{n} \delta_t$ for some non-negative m. Another example is the case of discounted payoffs, when $\theta = (\lambda(1-\lambda)^{t-1})_{t\geq 1}$ for some discount factor $\lambda \in (0, 1]$, in this case the value v_{θ} is denoted v_{λ} and $TV(\theta) = \lambda \xrightarrow[\lambda \to 0]{} 0$.

Definition 3.5. The Gambling House $\Gamma = (X, F, r)$ has a general limit value v^* if (v_{θ}) uniformly converges to v^* when $TV(\theta)$ goes to zero, i.e.:

$$\forall \varepsilon > 0, \exists \alpha > 0, \forall \theta, \ (TV(\theta) \le \alpha \implies (\forall x \in X, |v_{\theta}(x) - v^*(x)| \le \varepsilon)).$$

The existence of the general limit value implies in particular that $(v_n)_n$ and $(v_\lambda)_\lambda$ converge to the same limit when n goes to $+\infty$ and λ goes to 0. This is coherent with the result of Lehrer and Sorin [23], which states that the uniform convergence of $(v_n)_n$ and $(v_\lambda)_\lambda$ are equivalent. A recent characterization of the uniform convergence of a sequence of value functions $(v_{\theta^k})_k$, when $TV(\theta^k) \xrightarrow[k\to\infty]{} 0$, can be found in Renault [34], and it is shown that all such sequences have a unique possible limit point given by $v^* = \inf_{\theta \in \Delta(\mathbb{N}^*)} \sup_{m\geq 0} v_{m,\theta}$, where $v_{m,\theta}$ is the value corresponding to the evaluation with weight 0 for the first m stages and with weight θ_{t-m} for stages t > m.

In the definition of the general limit value, we require all value functions to be close to v^* when the patience is high, but the plays used may depend on the precise expression of θ . In the following definition, we require the same play to be simultaneously optimal for all θ patient enough.

Definition 3.6. The Gambling House $\Gamma = (X, F, r)$ has a general uniform value if it has a general limit value v^* and moreover for each $\varepsilon > 0$ one can find $\alpha > 0$ and for each initial state x a mixed play $\sigma(x)$ at x satisfying:

$$\forall \theta, \ (TV(\theta) \le \alpha \implies (\forall x \in X, \gamma_{\theta}(\sigma(x)) \ge v^*(x) - \varepsilon)).$$

The literature in repeated games has mainly focused on the evaluations $\theta = \sum_{t=1}^{n} \frac{1}{n} \delta_t$ and $\theta = (\lambda(1-\lambda)^{t-1})_{t\geq 1}$. The standard (Cesàro)-uniform value can be defined by restricting the evaluations to be Cesàro means: for each $\varepsilon > 0$ one can find n_0 and for each initial state x a mixed play $\sigma(x)$ at x satisfying: $\forall n \geq n_0, \forall x \in X, \gamma_n(\sigma(x)) \geq v^*(x) - \varepsilon$. Recently, Renault [32] considered deterministic Gambling Houses and characterized the uniform convergence of the value functions $(v_n)_n$. The existence of the standard Cesàro-uniform value is proved under some assumptions, including the case where the set of states X is metric precompact, the transitions are non expansive and the payoff function is uniformly continuous. As a corollary, the existence of the uniform value is shown in Partial Observation Markov Decision Processes with finite set of states (after each stage the decision-maker just observes a stochastic signal more or less correlated to the new state).

We now present our main theorem for Gambling Houses. Equation (1) implies that the general limit value v^* necessarily has to satisfy some rigidity property. The function v^* (or more precisely its linear extension to Z) can only be an "excessive function" in the terminology of potential theory [13] and gambling houses (Dubins and Savage [15], Maitra and Sudderth [24]).

Definition 3.7. An affine function w defined on Z (or $\Delta(X)$) is said to be excessive if for all x in X, $w(x) \ge \sup_{u \in F(x)} w(u)$.

Example 3.8. Let us consider the splitting transition given by a finite set K, $X = \Delta(K)$ and for each x in X, $F(x) = \{u \in \Delta(X), \sum_{p \in X} u(p) | p = x\}$ is the

set of probabilities on X centered at x. Then the function w from Z or $\Delta(X)$ to \mathbb{R} is excessive if and only if the restriction of w to X is concave. Moreover given $u, u' \in \Delta(X), u' \in \hat{F}(u)$ if and only if u' is the sweeping of u as defined by Choquet [13]: for all continuous concave functions f from X to $\mathbb{R}, u'(f) \leq u(f)$.

Assume now that X is a compact metric space and that r is continuous. r is naturally extended to an affine continuous function on $\Delta(X)$ by $r(u) = \int_{p \in X} r(p) du(p)$ for all Borel probabilities on X. In the following definition, we consider the closure of the graph of \hat{F} within the (compact) set $\Delta(X \times X)$.

Definition 3.9. An element u in $\Delta(X)$ is said to be an invariant measure of the Gambling House $\Gamma = (X, F, r)$ if $(u, u) \in cl(Graph \hat{F})$. The set of invariant measures of Γ is denoted by R, so that:

$$R = \{ u \in \Delta(X), (u, u) \in \operatorname{cl}(Graph \ \hat{F}) \}.$$

R is a convex compact subset of $\Delta(X)$. Even when Γ is deterministic, we still need to work in the space $\Delta(X)$ of probabilities over X to define invariant measures. Intuitively, we have replaced the time averages by the space averages. Recall that for u and u' in $\Delta(X)$, the Kantorovich-Rubinstein distance between u and u' is denoted by $d_{KR}(u, u') = \sup_{f \in E_1} |u(f) - u'(f)|$.

Theorem 3.10. Consider a Gambling House $\Gamma = (X, F, r)$ such that X is a compact metric space, r is continuous and F is non expansive with respect to the Kantorovich-Rubinstein distance:

$$\forall x \in X, \forall x' \in X, \forall u \in F(x), \exists u' \in F(x') \ s.t. \ d_{KR}(u, u') \le d(x, x').$$

Then the Gambling House Γ has a general uniform value v^* characterized by:

$$\forall x \in X, \ v^*(x) = \inf \left\{ w(x), w : \Delta(X) \to [0, 1] \text{ affine continuous s.t.} \\ (1) \ \forall y \in X, w(y) \ge \sup_{u \in F(y)} w(u) \text{ and } (2) \ \forall u \in R, w(u) \ge r(u) \right\}.$$

That is, v^* is the smallest continuous affine function on X which is 1) excessive and 2) above the running payoff r on invariant measures.

Notice that:

1) when $\Gamma = (X, F, r)$ is deterministic, the hypotheses are satisfied as soon as X is metric compact for some metric d, r is continuous and F is non expansive for d.

2) when X is finite, one can use the distance d(x, x') = 2 for all $x \neq x'$ in X, so that for u and u' in $\Delta(X)$, $d_{KR}(u, u') = ||u - u'||_1 = \sum_{x \in X} |u(x) - u'(x)|$, and the hypotheses are automatically satisfied. We will prove later a more general result for a model of MDP with finite state space, allowing for explicit actions influencing transitions and payoffs (see Corollary 3.21). **Remark 3.11.** The formula also holds when there is no decision-maker, i.e. when F is single-valued, and there are some similarities with the Von Neumann's ergodic theorem [42]. Let Z be a Hilbert space and Q be a linear isometry on Z, this theorem states that for all $z \in Z$, the sequence $z_n = \frac{1}{n} \sum_{t=1}^{n} Q^t(z)$ converges to the projection z^* of z on the set R of fixed points of Q. Using the linearity and the non expansiveness leads to a characterization by the set of fixed points. In particular, having in mind linear payoff functions of the form $(z \mapsto \langle l, z \rangle)$, we have that the projection z^* of z on R is characterized by:

 $\forall l \in Z, < l, z^* > = < l^*, z > = \inf\{< l', z >, l' \in R \text{ and } < l', r > \ge < l, r > \forall r \in R\}.$

Example 3.12. We consider here a basic periodic sequence of 0 and 1. Let $X = \{0, 1\}$ and for all $x \in X$, $F(x) = \{1 - x\}$ and r(x) = x. There is a unique invariant measure $u = 1/2\delta_0 + 1/2\delta_1$, and the general uniform value exists and satisifies $v^*(x) = \frac{1}{2}$ for all states x. Notice that considering evaluations $\theta = (\theta_t)_t$ such that θ_t is small for each t without requiring $TV(\theta)$ small, would not necessarily lead to v^* . Consider for instance $\theta^n = \sum_{t=1}^n \frac{1}{n} \delta_{2t}$ for each n, we have $v_{\theta^n}(x) = x$ for all x in X.

Example 3.13. The state space is the unit circle, let $X = \{x \in \mathbb{C}, |x| = 1\}$ and $F(e^{i\alpha}) = e^{i(\alpha+1)}$ for all real α . If we denote by μ the uniform distribution (Haar probability measure) on the circle, the mapping F is μ -ergodic and μ is F-invariant. By Birkhoff's theorem [8], we know that the time average converges to the space average μ -almost surely. Here μ is the unique invariant measure, and we obtain that the general uniform value is the constant:

$$\forall x \in X, \ v^*(x) = \frac{1}{2\pi} \int_0^{2\pi} r(e^{i\alpha}) d\alpha.$$

Notice that we obtain the convergence of the value $v_{\theta}(x)$ to $v^*(x)$ for all x in X, and not only for μ -almost all x in X.

Example 3.14. Let $\Gamma = (X, F, r)$ be a MDP satisfying the hypotheses of Theorem 3.10, and such that for all $x \in X$, $\delta_x \in F(x)$. Therefore the set R is equal to $\Delta(X)$. In the terminology of Gambling Theory (see Maitra Sudderth, [24], Γ is called a *leavable* gambling house since at each stage the player can stay at the current state. The limit value v^* is here characterized by:

$$v^* = \inf\{v : X \to [0, 1] \text{ continuous}, v \text{ is excessive } \text{ and } v \ge r\}.$$

In the above formula, v excessive means: $\forall x \in X, v(x) \geq \sup_{u \in F(x)} I\!\!E_u(v)$. This is a variant of the Fundamental Theorem of Gambling Theory (see section 3.1 in Maitra Sudderth [24]).

Example 3.15. The following deterministic Gambling House, which is an extension of example 1.4.4. in Sorin [41] and of example 5.2 of Renault [32],

shows that the assumptions of Theorem 3.10 allow for many speeds of convergence to the limit value v^* . Here l > 1 is a fixed parameter, X is the simplex $\{x = (p^a, p^b, p^c) \in \mathbb{R}^3_+, p^a + p^b + p^c = 1\}$ and the initial state is $x_0 = (1, 0, 0)$. The payoff is $r(p^a, p^b, p^c) = p^b - p^c$, and the transition is defined by: $F(p^a, p^b, p^c) = \{((1 - \alpha - \alpha^l)p^a, p^b + \alpha p^a, p^c + \alpha^l p^a), \alpha \in [0, 1/2]\}.$

The probabilistic interpretation is the following: there are 3 points a, b and c, and the initial point is a. The payoff is 0 at a, it is +1 at b, and -1 at c. At point a, the decision-maker has to choose $\alpha \in [0, 1/2]$: then b is reached with probability α , c is reached with probability α^l , and the play stays in a with the remaining probability $1 - \alpha - \alpha^l$. When b (resp. c) is reached, the play stays at b (resp. c) forever. So the decision-maker starting at point a wants to reach b and to avoid c. By playing at each stage $\alpha > 0$ small enough, he can get as close to b as he wants.

Back to our deterministic setup, we use norm $\|.\|_1$ and obtain that X is compact, F is non expansive and r is continuous, so that Theorem 3.10 applies, and the limit value is given by $v^*(p^a, p^b, p^c) = p^a + p^b$. Notice that even though the data are very smooth, there is no 0-optimal strategy here, in the sense that there is no mixed play σ at a such that $\lim_{TV(\theta)\to 0} \gamma_{\theta}(\sigma) = 1$, any good strategy requires to take a positive risk.

If we denote by x_{λ} the value $v_{\lambda}(x_0)$, we have for all $\lambda \in (0, 1]$: $x_{\lambda} = \phi_{\lambda}(x_{\lambda})$, where for all $x \in \mathbb{R}$,

$$\phi_{\lambda}(x) = \max_{\alpha \in [0, 1/2]} (1 - \lambda)(1 - \alpha - \alpha^{l})x + \alpha.$$

Since $x_{\lambda} \in (0, 1)$, the first order condition gives $(1 - \lambda)(-1 - l\alpha^{l-1})x_{\lambda} + 1 = 0$ and we can obtain:

$$x_{\lambda} = \frac{1}{(1-\lambda)} \left(l \left(\frac{\lambda}{(1-\lambda)(l-1)} \right)^{\frac{l-1}{l}} + 1 \right)^{-1}.$$

Finally we compute an equivalent of $x_{\lambda} - 1$ as λ goes to 0, and we obtain:

$$1 - v_{\lambda}(x_0) \sim C \lambda^{\frac{l-1}{l}}$$
 with $C = \frac{l}{(l-1)^{\frac{(l-1)}{l}}}$

3.2 Long-term values for standard MDPs

A standard Markov Decision Problem Ψ is given by a non empty set of states X, a non empty set of actions A, a mapping $q: X \times A \to \Delta_f(X)$ and a payoff function $g: X \times A \to [0, 1]$. At each stage, the player learns the current state x and chooses an action a. He then receives the payoff g(x, a), a new state is drawn accordingly to q(x, a) and the game proceeds to the next stage.

Definition 3.16. A pure, or deterministic, strategy is a sequence of mappings $\sigma = (\sigma_t)_{t\geq 1}$ where $\sigma_t : (X \times A)^{t-1} \to A$ for each t. A strategy (or behavioral strategy) is a sequence of mappings $\sigma = (\sigma_t)_{t\geq 1}$ where $\sigma_t : (X \times A)^{t-1} \to \Delta_f(A)$ for each t. We denote by Σ the set of strategies.

 $(X \times A)^0$ is a singleton, so σ_1 is viewed as an element of $\Delta_f(A)$ representing the lottery on actions played at the first stage (or simply if σ is pure, the action in A played at the first stage). A pure strategy is a particular case of strategy. An initial state x_1 in X and a strategy σ naturally induce a probability distribution with finite support over the set of finite histories $(X \times A)^n$ for all n, which can be uniquely extended to a probability over the set $(X \times A)^\infty$ of infinite histories.

Definition 3.17. Given an evaluation θ and an initial state x_1 in X, the θ -payoff of a strategy σ at x_1 is defined as $\gamma_{\theta}(x_1, \sigma) = \mathbb{E}_{x_1,\sigma} \left(\sum_{t \geq 1} \theta_t g(x_t, a_t) \right)$, and the θ -value at x_1 is:

$$v_{\theta}(x_1) = \sup_{\sigma \in \Sigma} \gamma_{\theta}(x_1, \sigma).$$

As for gambling houses, it is easy to see that the supremum can be taken over the smaller set of pure strategies, and one can derive a recursive formula linking the value functions. General limit and uniform values are defined as in subsection 3.1.

Definition 3.18. Let $\Psi = (X, A, q, g)$ be a standard MDP.

 Ψ has a general limit value v^* if (v_{θ}) uniformly converges to v^* when $TV(\theta)$ goes to zero, i.e. for each $\varepsilon > 0$ one can find $\alpha > 0$ such that:

$$\forall \theta, \ (TV(\theta) \le \alpha \implies (\forall x \in X, |v_{\theta}(x) - v^*(x)| \le \varepsilon)).$$

 Ψ has a general uniform value if it has a general limit value v^* , and if for each $\varepsilon > 0$ one can find $\alpha > 0$ and a behavior strategy $\sigma(x)$ for each initial state x satisfying:

$$\forall \theta, \ (TV(\theta) \le \alpha \implies (\forall x \in X, \gamma_{\theta}(x, \sigma(x)) \ge v^*(x) - \varepsilon)).$$

We now present a notion of invariance for the MDP Ψ . The next definition will be similar to Definition 3.9, however one needs to be slightly more sophisticated here to incorporate the payoff component. Assume now that X is a compact metric space, and define for each (u, y) in $\Delta_f(X) \times [0, 1]$,

$$\hat{F}(u,y) = \left\{ \left(\sum_{x \in X} u(x)q(x,a(x)), \sum_{x \in X} u(x)g(x,a(x)) \right), \text{ where } a: X \to \Delta_f(A) \right\}.$$

where q(x, .) and g(x, .) have been linearly extended for all x. We have defined a correspondence \hat{F} from $\Delta_f(X) \times [0, 1]$ to itself. It is easy to see that \hat{F} always is an affine correspondence (see Lemma 3.27 later). In the following definition we consider the closure of the graph of \hat{F} within the compact set $(\Delta(X) \times [0, 1])^2$, with the weak topology.

Definition 3.19. An element (u, y) in $\Delta(X) \times [0, 1]$ is said to be an invariant couple for the MDP Ψ if $((u, y), (u, y)) \in cl(Graph(\hat{F}))$. The set of invariant couples of Ψ is denoted by RR.

Our main result for standard MDPs is the following theorem, where X is assumed to be a compact subset of a simplex $\Delta(K)$, with K a finite set. Recall that $D_1 = \{f \in \mathcal{C}(\Delta(K)), \forall x, y \in \Delta(K), \forall a, b \ge 0, af(x) - bf(y) \le ||ax - by||_1\}$, and any f in D_1 is linearly extended to $\Delta(\Delta(K))$.

Theorem 3.20. Let $\Psi = (X, A, q, g)$ be a standard MDP where X is a compact subset of a simplex $\Delta(K)$, and such that:

$$\forall x \in X, \forall y \in X, \forall a \in A, \forall f \in D_1, \forall \alpha \ge 0, \forall \beta \ge 0,$$

 $|\alpha f(q(x,a)) - \beta f(q(y,a))| \le ||\alpha x - \beta y||_1$ and $|\alpha g(x,a) - \beta g(y,a)| \le ||\alpha x - \beta y||_1$. Then Ψ has a general uniform value v^* characterized by: for all x in X,

$$\begin{aligned} v^*(x) &= \inf \left\{ w(x), w : \Delta(X) \to [0, 1] \text{ affine continuous } s.t. \\ (1) \ \forall x' \in X, w(x') \geq \sup_{a \in A} w(q(x', a)) \text{ and } (2) \ \forall (u, y) \in RR, w(u) \geq y \right\}. \end{aligned}$$

The proof of Theorem 3.20 will be in section 3.4. An immediate corollary is when the state space is finite.

Corollary 3.21. Consider a standard MDP (K, A, q, g) with a finite set of states K. Then it has a general uniform value v^* , and for each state k:

$$\begin{aligned} v^*(k) &= \inf \left\{ w(k), w : \Delta(K) \to [0, 1] \text{ affine } s.t. \\ (1) \ \forall k' \in K, w(k') \geq \sup_{a \in A} w(q(k', a)) \text{ and } (2) \forall (p, y) \in RR, w(p) \geq y \right\}. \end{aligned}$$

with $RR = \{(p, y) \in \Delta(K) \times [0, 1], ((p, y), (p, y)) \in cl(conv(Graph(F)))\}$ and $F(k, y) = \{(q(k, a), g(k, a)), a \in A\}.$

Proof of Corollary 3.21: K is viewed as a subset of the simplex $\Delta(K)$, endowed with the L^1 -norm. Fix k, k' in K, a in A, $\alpha \ge 0$ and $\beta \ge 0$. We have

$$\|\alpha k - \beta k'\| = \begin{cases} |\alpha - \beta| \text{ if } k = k', \\ \alpha + \beta \text{ otherwise.} \end{cases}$$

First,

$$|\alpha g(k,a) - \beta g(k',a)| \le \begin{cases} |\alpha - \beta| g(k,a) \text{ if } k = k' \\ \alpha + \beta \text{ otherwise} \end{cases}$$

so in all cases $|\alpha g(k, a) - \beta g(k', a)| \leq ||\alpha k - \beta k'||$. Secondly, consider $f \in D_1$. f takes values in [-1, 1], so similarly we have: $|\alpha f(q(k, a)) - \beta f(q(k', a))| \leq ||\alpha k - \beta k'||$. So we can apply Theorem 3.20, and the graph of \hat{F} is the convex hull of the graph of F. **Remark 3.22.** When both sets of states and actions are finite, we are in the simple setting of Blackwell [9]. In this case our characterization leads to a dual formulation of a result of Denardo and Fox [14]. We say that a couple $(w, h) \in \mathbb{R}^K \times \mathbb{R}^K$ is superharmonic, in the sense of Hordjik and Kallenberg [20], if

$$\forall k \in K, \forall a \in A, \ w(k) + h(k) \ge g(k, a) + \sum_{k' \in K} q(k, a)(k')h(k').$$
 (2)

Denardo and Fox [14] showed that the value v^* is the smallest (pointwise) excessive function that can be completed by a function h such that (v^*, h) is superharmonic.

The existence of a function h such that (w, h) is superharmonic is equivalent to condition (2) of Corollary 3.21. Given a function w, solving equation (2) is a linear programming problem with $K \times A$ inequalities. By Farkas' lemma, it has a solution if and only if the following linear programming problem (D_w) , with unknown $\pi \in \mathbb{R}^{K \times A}$, has no solution:

$$\begin{array}{lll} \forall (k,a) \in K \times A & \pi(k,a) & \geq 0 \\ \forall k \in K & \sum_{a' \in A} \pi(k,a') & = \sum_{k' \in K, a' \in A} \pi(k',a')q(k',a')(k) \\ & \sum_{k' \in , a' \in A} \pi(k',a')g(k',a') & > \sum_{k' \in K, a' \in A} \pi(k',a')w(k'). \end{array}$$

Fix w a function on K. We prove that (D_w) has a solution if and only if condition (2) of Corollary 3.21 is not satisfied. Therefore both conditions are equivalent.

Let π be a solution of (D_w) . We can assume without loss of generality that $\sum_{k,a} \pi(k, a) = 1$. We denote by p the marginal of π on K and put for all $k \in K$, $\sigma(k) = \left(\frac{\pi(k,a')}{p(k)}\right)_{a' \in A} \in \Delta(A)$ if p(k) > 0, and define arbitrarily $\sigma(k)$ if p(k) = 0. Denote by σ the strategy which plays $\sigma(k)$ if the state is k for each k. The payoff obtained by playing σ from distribution p is $y = \sum_{k' \in a' \in A} \pi(k', a')g(k', a')$. The second line of equations of (D_w) implies that p is invariant by σ , so $(p, y) \in RR$ and the last equation implies that y > w(p). The function w does not satisfy condition (2) of Corollary 3.21.

Conversally if condition (2) of Corollary 3.21 is not satisfied, there exists $(p, y) \in RR$ such that y > w(p). By compacity of the set of probabilities over K and the set of payoffs, there exists a strategy $\sigma \in \Delta(A)^K$, such that p is invariant under σ and y is the payoff obtained by playing σ from distribution p. The probability π defined by

$$\forall k \in K, \forall a \in A, \pi(k, a) = p(k)\sigma(k, a).$$

is then a solution of (D_w) .

Note that Denardo and Fox also use duality theory but they study directly the minimization problem with unknown w and h and deduce a dual maximization problem.

3.3 Proof of Theorem 3.10

In this section we consider a compact metric space (X, d), and we use the Kantorovich-Rubinstein distance $d = d_{KR}$ on $\Delta(X)$. We write $Z = \Delta_f(X)$, $\overline{Z} = \Delta(X)$. We start with a lemma.

Lemma 3.23. Let $F : X \rightrightarrows \Delta_f(X)$ be non expansive for d_{KR} . Then the mixed extension of F is 1-Lipschitz from $\Delta_f(X)$ to $\Delta_f(X)$ for d_{KR} .

Proof of Lemma 3.23. We first show that the mapping $(p \mapsto \operatorname{conv} F(p))$ is non expansive from X to Z. Indeed, consider p and p' in X, and $u = \sum_{i \in I} \alpha_i u_i$, with I finite, $\alpha_i \geq 0$, $u_i \in F(p)$ for each i, and $\sum_{i \in I} \alpha_i = 1$. By assumption for each i one can find u'_i in F(p') such that $d_{KR}(u_i, u'_i) \leq d(p, p')$. Define $u' = \sum_{i \in I} \alpha_i u'_i$ in $\operatorname{conv} F(p')$. We have:

$$d_{KR}(u, u') = \sup_{f \in E_1} \left(\sum_i \alpha_i u_i(f) - \sum_i \alpha_i u'_i(f) \right),$$

$$= \sup_{f \in E_1} \sum_{i \in I} \alpha_i (u_i(f) - u'_i(f)),$$

$$\leq \sum_{i \in I} \alpha_i d_{KR}(u_i, u'_i),$$

$$\leq d(p, p').$$

We now prove that \hat{F} is 1-Lipschitz from Z to Z. Let u_1 , u_2 be in Z and $v_1 = \sum_{p \in X} u_1(p) f_1(p)$, where $f_1(p) \in \operatorname{conv} F(p)$ for each p. By the Kantorovich duality formula, there exists a coupling $\gamma = (\gamma(p,q))_{(p,q)\in X\times X}$ in $\Delta_f(X\times X)$ with first marginal u_1 and second marginal u_2 satisfying:

$$d_{KR}(u_1, u_2) = \sum_{(p,q) \in X \times X} \gamma(p,q) d(p,q).$$

For each p, q in X by the first part of this proof there exists $f^p(q) \in \operatorname{conv} F(q)$ such that $d_{KR}(f^p(q), f_1(p)) \leq d(p, q)$. We define:

$$f_2(q) = \sum_{p \in X} \frac{\gamma(p, q)}{u_2(q)} f^p(q) \in \operatorname{conv} F(q), \text{ and } v_2 = \sum_{q \in X} u_2(q) f_2(q) \in \hat{F}(u_2).$$

We now conclude.

$$d_{KR}(v_{1}, v_{2}) = d_{KR} \left(\sum_{p \in X} u_{1}(p) f_{1}(p), \sum_{q \in X} u_{2}(q) f_{2}(q) \right)$$

$$= d_{KR} \left(\sum_{p,q} \gamma(p,q) f_{1}(p), \sum_{q,p} \gamma(p,q) f^{p}(q) \right)$$

$$\leq \sum_{p,q} \gamma(p,q) d_{KR}(f_{1}(p), f^{p}(q))$$

$$\leq \sum_{p,q} \gamma(p,q) d(p,q) = d_{KR}(u_{1}, u_{2}).$$

The mixed extension of F is 1-lipschitz.

We now consider a Gambling House $\Gamma = (X, F, r)$ and assume the hypotheses of Theorem 3.10 are satisfied. We will work² with the deterministic Gambling House $\hat{\Gamma} = (\Delta_f(X), \hat{F}, r)$. Recall that r is extended to an affine and continuous mapping on $\Delta(X)$ whereas \hat{F} is an affine non expansive correspondence from Zto Z.

For p in X, the pure plays in $\hat{\Gamma}$ at the initial state δ_p coincide with the mixed plays in Γ at the initial state p. As a consequence, the θ -value for Γ at p coincides with the θ -value for $\hat{\Gamma}$ at δ_p , which is written $v_{\theta}(p) = v_{\theta}(\delta_p)$. Because \hat{F} and r are affine on Z, the θ -value for $\hat{\Gamma}$, as a function defined on Z, is the affine extension of the original v_{θ} defined on X. So we have a unique value function v_{θ} which is defined on Z and is affine. Because \hat{F} is 1-Lipschitz and r is uniformly continuous, all the value functions v_{θ} have the same modulus of continuity as r, so $(v_{\theta})_{\theta}$ is an equicontinuous family of mappings from Z to [0, 1]. Consequently, we extend v_{θ} to an affine mapping on \overline{Z} with the same modulous of continuity, and the family $(v_{\theta})_{\theta}$ now is an equicontinuous³ family of mappings from \overline{Z} to [0, 1].

We define R and v^* as in the statements of Theorem 3.10, so that for all x in X,

$$v^*(x) = \inf \{w(x), w : \overline{Z} \to [0, 1] \text{ affine continuous s.t.}$$

$$(1) \ \forall y \in X, w(y) \ge \sup_{u \in F(y)} w(u) \text{ and } (2) \forall u \in R, w(u) \ge r(u) \}.$$

We start with a lemma based on the non-expansiveness of \hat{F} .

Lemma 3.24. 1) Given (u, u') in $cl(Graph(\hat{F}))$, v in Z and $\varepsilon > 0$, there exists $v' \in \hat{F}(v)$ such that $d(u', v') \leq d(u, v) + \varepsilon$.

2) Given a sequence $(z_t)_{t\geq 0}$ of elements of \overline{Z} such that $(z_t, z_{t+1}) \in cl(Graph(\hat{F}))$ for all $t \geq 1$, and given $\varepsilon > 0$, one can find a sequence $(z'_t)_{t\geq 0}$ of elements of Zsuch that $(z'_t)_{t\geq 1}$ is a play at z'_0 , and $d(z_t, z'_t) \leq \varepsilon$ for each $t \geq 0$.

Proof of Lemma 3.24: 1) For all $\varepsilon > 0$ there exists $(z, z') \in Graph(\bar{F})$ such that $d(z, u) \leq \varepsilon$ and $d(z', u') \leq \varepsilon$. Because \hat{F} is non expansive, one can find v' in $\hat{F}(v)$ such that $d(z', v') \leq d(z, v)$. Consequently, $d(v', u') \leq d(v', z') + d(z', u') \leq d(z, v) + \varepsilon \leq d(u, v) + 2\varepsilon$.

2) It is first easy to construct (z'_0, z'_1) in the graph of \hat{F} such that $d(z'_0, z_0) \leq \varepsilon$ and $d(z'_1, z_1) \leq \varepsilon$. $(z_1, z_2) \in cl(Graph(\hat{F}))$ so by 1) one can find (z'_2) in $\hat{F}(z'_1)$

²A variant of the proof would be to consider the Gambling House on $\Delta(X)$ where the transition correspondence is defined so that its graph is the closure of the graph of \hat{F} . Part 1) of Lemma 3.24 shows this correspondence is also non expansive.

 $^{{}^{3}}Z$ being precompact, this is enough to obtain the existence of a general limit value, see Renault [34]. Here we will moreover obtain a characterization of this value and the existence of the general uniform value.

such that $d(z_2, z'_2) \leq d(z_1, z'_1) + \varepsilon^2 \leq \varepsilon + \varepsilon^2$. Iterating, we construct a play $(z'_t)_{t\geq 1}$ at z'_0 such that $d(z_t, z'_t) \leq \varepsilon + \varepsilon^2 + \ldots + \varepsilon^t$ for each t.

Proposition 3.25. Γ has a general limit value given by v^* .

Proof of Proposition 3.25: By Ascoli's theorem, it is enough to show that any limit point of $(v_{\theta})_{\theta}$ (for the uniform convergence) coincides with v^* . We thus assume that $(v_{\theta^k})_k$ uniformly converges to v on \overline{Z} when k goes to ∞ , for a family of evaluations satisfying:

$$\sum_{t\geq 1} |\theta_{t+1}^k - \theta_t^k| \longrightarrow_{k\to\infty} 0.$$

And we need to show that $v = v^*$.

A) We first show that $v \ge v^*$.

It is plain that v can be extended to an affine function on \overline{Z} and has the modulus of continuity of r. Because $\sum_{t\geq 1} |\theta_{t+1}^k - \theta_t^k| \longrightarrow_{k\to\infty} 0$, we have by Equation (1) of section 3.1 that: $\forall y \in X, v(y) = \sup_{u \in F(y)} v(u)$.

Let now u be in R. By Lemma 3.24 for each ε one can find u_0 in Z and a play $(u_1, u_2, ..., u_t, ...)$ such that $u_t \in \hat{F}(u_{t-1})$ and $d(u, u_t) \leq \varepsilon$ for all $t \geq 0$. Because r is uniformly continuous, we get $v(u) \geq r(u)$.

By definition of v^* as an infimum, we obtain: $v^* \leq v$.

B) We show that $v^* \geq v$. Let w be a continous affine mapping from \overline{Z} to [0,1] satisfying (1) and (2) of the definition of v^* . It is enough to show that $w(p) \geq v(p)$ for each p in X. Fix p in X and $\varepsilon > 0$.

For each k, let $\sigma^k = (u_1^k, ..., u_t^k, ...) \in Z^{\infty}$ be a play at δ_p for $\hat{\Gamma}$ which is almost optimal for the θ^k -value, in the sense that $\sum_{t\geq 1} \theta_t^k r(u_t^k) \geq v_{\theta^k}(p) - \varepsilon$. Define:

$$u(k) = \sum_{t=1}^{\infty} \theta_t^k u_t^k \in \overline{Z}, \text{ and } u'(k) = \sum_{t=1}^{\infty} \theta_t^k u_{t+1}^k \in \overline{Z}.$$

u(k) and u'(k) are well-defined limits of normal convergent series in the Banach space $\mathcal{C}(X)'$. Because \hat{F} is affine, its graph is a convex set and $(u(k), u'(k)) \in$ $\mathrm{cl}(Graph(\hat{F}))$ for each k.

Moreover, we have $d(u(k), u'(k)) \leq \operatorname{diam}(X)(\theta_1^k + \sum_{t=2}^{\infty} |\theta_t^k - \theta_{t-1}^k|)$, where $\operatorname{diam}(X)$ is the diameter of X. Consequently, $\sum_{t\geq 1} |\theta_{t+1}^k - \theta_t^k| \longrightarrow_{k\to\infty} 0$ implies $d(u(k), u'(k)) \longrightarrow_{k\to\infty} 0$. Considering a limit point of the sequence $(u(k), u'(k))_k$, we obtain some u in R. By assumption on w, $w(u) \geq r(u)$. Moreover, for each k we have $r(u(k)) = \sum_{t\geq 1} \theta_t^k r(u_t^k) \geq v_{\theta^k}(p) - \varepsilon$, so $r(u) \geq v(p) - \varepsilon$.

Because w is excessive, we obtain that for each k the sequence $(w(u_t^k))_t$ is non increasing, so $w(u(k)) = \sum_{t>1} \theta_t^k w(u_t^k) \le w(p)$. So we obtain:

$$w(p) \ge w(u) \ge r(u) \ge v(p) - \varepsilon.$$

This is true for all ε , so $w \ge v$.

Proposition 3.26. Γ has a general uniform value.

Proof of Proposition 3.26: First we can extend the notion of mixed play to Z. A mixed play at $u_0 \in Z$, is a sequence $\sigma = (u_1, ..., u_t, ...) \in Z^{\infty}$ such that $u_{t+1} \in \hat{F}(u_t)$ for each $t \geq 0$, and we denote by $\Sigma(u_0)$ the set of mixed plays at u_0 . Given t, T in \mathbb{N} , $n \in \mathbb{N}^*$ and $u_0 \in Z$, we define for each mixed play $\sigma = (u_t)_{t\geq 1} \in \Sigma(u_0)$ the auxiliary payoff:

$$\gamma_{t,n}(\sigma) = \frac{1}{n} \sum_{l=t+1}^{t+n} r(u_l), \text{ and } \beta_{T,n}(\sigma) = \inf_{t \in \{0,\dots,T\}} \gamma_{t,n}(\sigma).$$

And we also define the auxiliary value function: for all u in Z,

$$h_{T,n}(u_0) = \sup_{\sigma \in \Sigma(u_0)} \beta_{T,n}(\sigma).$$

Clearly, $\beta_{T,n}(\sigma) \leq \gamma_{0,n}(\sigma)$ and $h_{T,n}(u_0) \leq v_n(u_0)$. We can write:

$$h_{T,n}(u_0) = \sup_{\sigma \in \Sigma(u_0)} \inf_{\theta \in \Delta(\{0,...,T\})} \frac{1}{n} \sum_{t=0}^T \theta_t \sum_{l=t+1}^{t+n} r(u_l)$$

=
$$\sup_{\sigma \in \Sigma(u_0)} \inf_{\theta \in \Delta(\{0,...,T\})} \sum_{l=1}^{T+n} \beta_l(\theta, n) r(u_l).$$

where for each l in 1, ..., T + n,

$$\beta_l(\theta, n) = \frac{1}{n} \sum_{t=Max\{0, l-n\}}^{\operatorname{Min}\{T, l-1\}} \theta_t.$$

By construction, \hat{F} is affine, so $\Sigma(u_0)$ is a convex subset of Z^{∞} . $\Delta(\{0, ..., T\})$ is convex compact and the payoff $\sum_{l=1}^{T+n} \beta_l(\theta, n) r(u_l)$ is affine both in θ and in σ . We can apply a standard minmax theorem to get:

$$h_{T,n}(u_0) = \inf_{\theta \in \Delta(\{0,\dots,T\})} \sup_{\sigma \in \Sigma(u_0)} \sum_{l=1}^{T+n} \beta_l(\theta, n) r(u_l).$$

We write $\theta_t = 0$ for t > T and for each $l \ge 0$: $\beta_l(n, \theta) = \frac{1}{n}(\theta_0 + ... + \theta_{l-1})$ if $l \le n$, $\beta_l(\theta, n) = \frac{1}{n}(\theta_{l-n} + ... + \theta_{l-1})$ if $n + 1 \le l \le n + T$, $\beta_l(n, \theta) = 0$ if l > n + T. The evaluation $\beta(\theta, n)$ is a particular probability on stages and $h_{T,n}(u_0) = \inf_{\theta \in \Delta(\{0,...,T\})} v_{\beta(\theta,n)}(u_0)$. It is easy to bound the total variation of $\beta(\theta, n)$:

$$\sum_{l\geq 0} |\beta_{l+1}(\theta, n) - \beta_l(\theta, n)| = \sum_{l=0}^{n-1} \frac{\theta_l}{n} + \sum_{l\geq n} \frac{1}{n} |\theta_l - \theta_{l-n}| \le \frac{3}{n} \longrightarrow_{n \to \infty} 0.$$

The impatience of $\beta(\theta, n)$ goes to zero as n goes to infinity, uniformly in θ . So we can use the previous Proposition 3.25 to get:

$$\forall \varepsilon > 0, \exists n_0, \forall n \ge n_0, \forall \theta \in \Delta(\mathbb{I} \mathbb{N}), \forall u_0 \in \mathbb{Z}, \ |v_{\beta(\theta, n)}(u_0) - v^*(u_0)| \le \varepsilon.$$

This implies that $h_{\infty,n}(u_0) :=_{def} \inf_{\theta \in \Delta(\mathbb{N})} v_{\beta(\theta,n)}(u_0) = \inf_{T \ge 0} h_{T,n}(u_0)$ converges to $v^*(u_0)$ when $n \to \infty$, and the convergence is uniform over Z. Consequently, if we fix $\varepsilon > 0$ there exists n_0 such that for all u_0 in Z, for all $T \ge 0$, there exists a play $\sigma^T = (u_t^T)_{t \ge 1}$ in $\Sigma(u_0)$ such that the average payoff is good on every interval of n_0 stages starting before T+1: for all $t = 0, ..., T, \gamma_{t,n_0}(\sigma^T) \ge v^*(u_0) - \varepsilon$.

We fix u_0 in Z and consider, for each T, the play $\sigma^T = (u_t^T)_{t\geq 1}$ in $\Sigma(u)$ as above. By a diagonal argument we can construct for each $t \geq 1$ a limit point z_t in \overline{Z} of the sequence $(u_t^T)_{T\geq 0}$ such that for all t we have $(z_t, z_{t+1}) \in cl(Graph(\hat{F}))$, with $z_0 = u_0$. For each $m \geq 0$, we have $\frac{1}{n_0} \sum_{t=m+1}^{m+1+n_0} r(u_t^T) \geq v^*(u_0) - \varepsilon$ for T large enough, so at the limit we get: $\frac{1}{n_0} \sum_{t=m+1}^{m+1+n_0} r(z_t) \geq v^*(u_0) - \varepsilon$. r being uniformly continuous, there exists α such that $|r(z) - r(z')| \leq \varepsilon$ as soon

r being uniformly continuous, there exists α such that $|r(z) - r(z')| \leq \varepsilon$ as soon as $d(z, z') \leq \alpha$. By Lemma 3.24, one can find a mixed play $\sigma' = (z'_1, ..., z'_t, ...)$ at $\Sigma(z_0)$ such that for each t, $d(z_t, z'_t) \leq \alpha$. We obtain that for each $m \geq 0$, $\frac{1}{n_0} \sum_{t=m+1}^{m+1+n_0} r(z'_t) \geq v^*(u) - 2\varepsilon$.

Consequently we have proved: $\forall \varepsilon > 0$, there exists n_0 such that for all initial state p in X, there exists a mixed play $\sigma' = (z'_t)_t$ at p such that: $\forall m \ge 0$, $\frac{1}{n_0} \sum_{t=m+1}^{m+1+n_0} r(z'_t) \ge v^*(p) - 2\varepsilon$.

Let $\theta \in \Delta(\mathbb{N}^*)$ be an evaluation, it is now easy to conclude. First if $v^*(p) - 2\epsilon < 0$, then any play is 2ϵ -optimal. Otherwise, for each $j \ge 1$, denote by $\overline{\theta_j}$ the maximum of θ on the block $B^j = \{(j-1)n_0 + 1, ..., jn_0\}$. For all $t \in B^j$, we have:

$$\overline{\theta_j} \ge \theta_t \ge \overline{\theta_j} - \sum_{t' \in \{(j-1)n_0 + 1, \dots, jn_0 - 1\}} |\theta_{t'+1} - \theta_{t'}|.$$

As a consequence, for all j we have:

$$\sum_{t=(j-1)n_0+1}^{jn_0} \theta_t r(z'_t) \ge \overline{\theta_j} \sum_{t=(j-1)n_0+1}^{jn_0} r(z'_t) - n_0 \sum_{t' \in \{(j-1)n_0+1,\dots,jn_0-1\}} |\theta_{t'+1} - \theta_{t'}|$$
$$\ge \sum_{t=(j-1)n_0+1}^{jn_0} \theta_t (v^*(p) - 2\varepsilon) - n_0 \sum_{t' \in \{(j-1)n_0+1,\dots,jn_0-1\}} |\theta_{t'+1} - \theta_{t'}|$$

and by summing over j, we get: $\gamma_{\theta}(x_0, \sigma') \ge v^*(p) - 2\epsilon - n_0 TV(\theta) \ge v^*(p) - 3\epsilon$ as soon as $TV(\theta)$ is small enough. \Box

3.4 Proof of Theorem 3.20

Assume that X is a compact subset of a simplex $\Delta(K)$, and let $\Psi = (X, A, q, g)$ be a standard MDP such that: $\forall x \in X, \forall y \in X, \forall a \in A, \forall f \in D_1, \forall a \ge 0, \forall \beta \ge 0$,

$$|\alpha f(q(x,a)) - \beta f(q(y,a))| \le \|\alpha x - \beta y\|_1 \text{ and } |\alpha g(x,a) - \beta g(y,a)| \le \|\alpha x - \beta y\|_1$$

We write $Z = \Delta_f(X) \times [0, 1]$, and $\overline{Z} = \Delta(X) \times [0, 1]$. We will use the metric $d_* = d_0 = d_1 = d_2 = d_3$ on $\Delta(\Delta(K))$ introduced in section 2.3 and its restriction to $\Delta(X)$, so that \overline{Z} is a compact metric space. For all $(u, y), (u', y') \in \Delta_f(X) \times [0, 1]$, we put $d((u, y), (u', y')) = \max(d_*(u, u'), |y - y'|)$ so that (Z, d) is a precompact metric space. Recall we have defined the correspondence \hat{F} from Z to itself such that for all (u, y) in Z,

$$\hat{F}(u,y) = \{ (Q(u,\sigma), G(u,\sigma)) \text{ s.t. } \sigma : X \to \Delta_f(A) \},\$$

with the notations $Q(u, \sigma) = \sum_{x \in X} u(x)q(x, \sigma(x))$ and $G(u, \sigma) = \sum_{x \in X} u(x)g(x, \sigma(x))$. And we simply define the payoff function r from Z to [0, 1] by r(u, y) = y for all (u, y) in Z. We start with a crucial lemma, which shows the importance of the duality formula of Theorem 2.19.

Lemma 3.27. \hat{F} is an affine and non expansive correspondence from Z to itself.

Proof of Lemma 3.27. We first show that: $\forall u, u' \in \Delta_f(X), \forall \alpha \in [0, 1], \forall y, y' \in [0, 1], \hat{F}(\alpha u + (1 - \alpha)u', \alpha y + (1 - \alpha)y') = \alpha \hat{F}(u, y) + (1 - \alpha)\hat{F}(u', y').$ First the transition does not depend on the second coordinate so we can forget it for the rest of the proof. The \subset part is clear. To see the reverse inclusion, consider $\sigma : X \to \Delta_f(A), \sigma' : X \to \Delta_f(A)$ and $v = \alpha \sum_{x \in X} u(x)q(x, \sigma(x)) + (1 - \alpha)\sum_{x \in X} u'(x)q(x, \sigma'(x))$ in $\alpha \hat{F}(u) + (1 - \alpha)\hat{F}(u')$. Define

$$\sigma^*(x) = \frac{\alpha u(x)\sigma(x) + (1-\alpha)u'(x)\sigma'(x)}{\alpha u(x) + (1-\alpha)u'(x)}$$

for each x such that the denominator is positive. Then $v = \sum_{x \in X} (\alpha u + (1 - \alpha)u'(x))q(x, \sigma^*(x))$, and \hat{F} is affine.

We now prove that \hat{F} is non expansive. Let z = (u, y) and z' = (u', y')be in Z. We have $d((u, y), (u', y')) \geq d_*(u, u')$ and denote by U and U' the respective supports of u and u'. By the duality formula of Theorem 2.19, there exists $\alpha = (\alpha(p, p'))_{(p,p') \in U \times U'}$ and $\beta = (\beta(p, p'))_{(p,p') \in U \times U'}$ with non-negative coordinates satisfying: $\sum_{p' \in U'} \alpha(p, p') = u(p)$ for all $p \in U$, $\sum_{p \in U} \beta(p, p') = u'(p')$ for all $p' \in U'$, and

$$d_*(u, u') = \sum_{(p, p') \in U \times U'} \| p \; \alpha(p, p') - p' \; \beta(p, p') \|_1.$$

Consider now $v = Q(u, \sigma) = \sum_{p \in U} u(p)q(p, \sigma(p))$ for some $\sigma : X \to \Delta_f(A)$. We define for all p' in U':

$$\sigma'(p') = \sum_{p \in U} \frac{\beta(p, p')}{u'(p')} \sigma(p),$$

and $v' = Q(u', \sigma') = \sum_{p' \in U'} u'(p')q(p', \sigma'(p'))$. Then $v' \in \hat{F}(u', y')$, and for each test function φ in D_1 we have:

$$\begin{aligned} |\varphi(v) - \varphi(v')| &= |\sum_{p,p'} \alpha(p,p')\varphi(q(p,\sigma(p))) - \beta(p,p')\varphi(q(p',\sigma(p)))| \\ &= |\sum_{p,p',a} \alpha(p,p')\sigma(p)(a)\varphi(q(p,a)) - \beta(p,p')\sigma(p)(a)\varphi(q(p',a))| \\ &\leq \sum_{p,p'} ||\alpha(p,p')p - \beta(p,p')p'||_1 = d_*(u,u'), \end{aligned}$$

and therefore $d_*(v, v') \leq d_*(u, u')$. In addition we have a similar result on the payoff,

$$|G(u,\sigma) - G(u',\sigma')| = |\sum_{p,p'} \alpha(p,p')g(p,\sigma(p)) - \beta(p,p')g(p',\sigma(p))|$$

$$\leq \sum_{p,p'} ||\alpha(p,p')p - \beta(p,p')p'||_1$$

$$\leq d_*(u,u').$$

Thus we have $d((Q(u, \sigma), R(u, \sigma)), (Q(u', \sigma'), R(u', \sigma'))) \le d_*(u, u') \le d(z, z')$. \Box

Recall that the set of invariant couples of the MDP Ψ is:

$$RR = \{(u, y) \in \overline{Z}, ((u, y), (u, y)) \in cl(Graph(\hat{F}))\},\$$

and the function $v^*: X \longrightarrow \mathbb{R}$ is defined by:

$$v^*(x) = \inf \{ w(x), w : \Delta(X) \to [0, 1] \text{ affine } continuous \ s.t.$$

$$(1) \ \forall y \in X, w(y) \ge \sup_{a \in A} w(q(y, a)) \text{ and } (2) \ \forall (u, y) \in RR, w(u) \ge y \}.$$

We now consider the deterministic Gambling House $\hat{\Gamma} = (Z, \hat{F}, r)$. Z is precompact metric, \hat{F} is affine non expansive and r is obviously affine and uniformly continuous. Given an evaluation θ , the θ -value of $\hat{\Gamma}$ at $z_0 = (u, y)$ is denoted by $\hat{v}_{\theta}(u, y) = \hat{v}_{\theta}(u)$ and does not depend on y. The recursive formula of section 3.1 yields:

$$\begin{aligned} \forall (u,y) \in Z, \quad \hat{v}_{\theta}(u) &= \sup_{\substack{(u',y') \in \hat{F}(u) \\ \sigma \in X \to \Delta_f(A)}} \theta_1 y' + (1-\theta_1) \hat{v}_{\theta^+}(u') \\ &= \sup_{\sigma \in X \to \Delta_f(A)} \left(\theta_1 G(u,\sigma) + (1-\theta_1) \hat{v}_{\theta^+}(Q(u,\sigma)) \right). \end{aligned}$$

Because \hat{F} and r are affine, \hat{v}_{θ} is affine in u and the supremum in the above expression can be taken over functions from X to A. Because \hat{F} is non expansive and r is 1-Lipschitz, each \hat{v}_{θ} is 1-Lipschitz.

We denote by v_{θ} the θ -value of the MDP Ψ and linearly extend it to $\Delta_f(X)$. It turns out that the recursive formula satisfied by v_{θ} is similar to the above recursive formula for \hat{v}_{θ} , so that $v_{\theta}(u) = \hat{v}_{\theta}(u, y)$ for all u in $\Delta_f(X)$ and y in [0, 1]. As a consequence, the existence of the general limit value in both problems $\hat{\Gamma}$ and Ψ is equivalent. Moreover, a deterministic play in $\hat{\Gamma}$ induces a strategy in Ψ , so that the existence of the general uniform value in $\hat{\Gamma}$ will imply the existence of the general uniform value in Ψ (note that deterministic and mixed plays in $\hat{\Gamma}$ are equivalent since \hat{F} has convex values).

It is thus sufficient to show that $\hat{\Gamma}$ has a general uniform value given by v^* , and we can mimic the end of the proof of Theorem 3.10. Lemma 3.24 applies word for word. Finally, one can proceed almost exactly as in Propositions 3.25 and 3.26 to show that $\hat{\Gamma}$, hence Ψ , has a general uniform value given by v^* .

4 Applications to partial observation and games

4.1 **POMDP** with finitely many states

We now consider a more general model of MDP with actions where after each stage, the decision-maker does not perfectly observe the state. A MDP with partial observation, or POMDP, $\Gamma = (K, A, S, q, g)$ includes a finite set of states K, a non empty set of actions A and a non empty set of signals S. The transition q now goes from $K \times A$ to $\Delta_f(S \times K)$ and the payoff function g still goes from $K \times A$ to [0, 1]. Given an initial probability p_1 on K, the POMDP $\Gamma(p_1)$ is played as follows. An initial state k_1 in K is selected according to p_1 and is not told to the decision-maker. At every stage $t \geq 1$, the decision-maker selects an action $a_t \in A$. If the current state is k_t , he has a (unobserved) payoff $g(k_t, a_t)$ and a pair (s_t, k_{t+1}) is drawn according to $q(k_t, a_t)$. Then the player learns s_t , and the play proceeds to stage t + 1 with new state k_{t+1} . A behavioral strategy is now a sequence $(\sigma_t)_{t\geq 1}$ of applications with for each t, $\sigma_t : (A \times S)^{t-1} \to \Delta_f(A)$. As usual, an initial probability on K and a behavior strategy σ induce a probability distribution over $(K \times A \times S)^{\infty}$ and we can define the θ -values and the notions of general limit and uniform values accordingly.

Theorem 4.1. A POMDP with finitely many states has a general uniform value, i.e. there exists $v^* : \Delta(K) \to \mathbb{R}$ with the following property: for each $\varepsilon > 0$ one can find $\alpha > 0$ and for each initial probability p a behavior strategy $\sigma(p)$ such that for each evaluation θ with $TV(\theta) \leq \alpha$,

 $\forall p \in \Delta(K), |v_{\theta}(p) - v^*(p)| \le \varepsilon \text{ and } \gamma_{\theta}(\sigma(p)) \ge v^*(p) - \varepsilon.$

Proof: It is natural to introduce an auxiliary MDP with state variable the belief of the decision-maker on the state in K. We define Ψ the standard MDP on

 $X = \Delta(K)$ with the same set of actions A and the following payoff and transition functions:

- $r: X \times A \longrightarrow [0, 1]$ s.t. $r(p, a) = \sum_{k \in K} p(k)g(k, a)$ for all p, a,
- $\hat{q}: X \times A \to \Delta_f(X)$ such that

$$\hat{q}(p,a) = \sum_{s \in S} \left(\sum_{k} p^{k} q(k,a)(s) \right) \delta_{\hat{q}(p,a,s)},$$

where $\hat{q}(p, a, s) \in \Delta(K)$ is the belief on the new state after playing a at p and observing the signal s:

$$\forall k' \in K, \hat{q}(p, a, s)(k') = \frac{q(p, a)(k', s)}{q(p, a)(s)} = \frac{\sum_{k} p^{k} q(k, a)(k', s)}{\sum_{k} p^{k} q(k, a)(s)}.$$

The POMDP $\Gamma(p_1)$ and the standard MDP $\Psi(p_1)$ have the same value for all θ -evaluations. And for each strategy σ in $\Psi(p_1)$, the player can guarantee the same payoff in the original game $\Gamma(p_1)$ by mimicking the strategy σ . So if we prove that Ψ has a general uniform value it will imply that the POMDP Γ has a general uniform value.

To conclude the proof, we will simply apply Theorem 3.20 to the MDP Ψ . We need to check the assumptions on the payoff and on the transition.

Consider any p, p' in $X, a \in A, \alpha \ge 0$ and $\beta \ge 0$. We have:

$$|\alpha r(p,a) - \beta r(p',a)| = \left| \sum_{k} (\alpha p(k) - \beta p'(k))g(k,a) \right| \le ||\alpha p - \beta p'||$$

Moreover for any $f \in D_1$, we have:

$$\begin{aligned} |\alpha \hat{q}(p,a)(f) - \beta \hat{q}(p',a)(f)| &= \left| \sum_{s \in S} \left(\alpha q(p,a)(s) f(\hat{q}(p,a,s)) - \beta q(p',a)(s) f(\hat{q}(p',a,s)) \right) \right| \\ &\leq \sum_{s} ||\alpha q(p,a)(.,s) - \beta q(p',a)(.,s)|| \\ &\leq \sum_{s,k,k'} |\alpha p(k')q(k',a)(k,s) - \beta p'(k')q(k',a)(k,s)| \\ &\leq \sum_{s,k,k'} q(k',a)(k,s) |\alpha p(k') - \beta p'(k')| = ||\alpha p - \beta p'||. \end{aligned}$$

where the first inequality comes from the definition of D_1 .

By Theorem 3.20, the MDP Ψ has a general uniform value and we deduce that the POMDP Γ has a general uniform value.

Example 4.2. Let $\Gamma = (K, A, S, q, g, p_1)$ be a POMDP where $K = \{k_1, k_2\}$, $A = \{a, b\}, S = \{*\}$ and $p_1 = \delta_{k_1}$. The initial state is k_1 and since there is only one

signal, the decision-maker will obtain no additional information on the state. We say that he is in the dark. The payoff is given by $g(k_1, a) = g(k_1, b) = g(k_2, b) = 0$ and $g(k_2, a) = 1$, and the transition by $q(k_1, a) = q(k_1, b) = \delta_{*,k_1}$, $q(k_2, a) = \delta_{*,k_2}$ and $q(k_1, b) = \frac{1}{2}\delta_{*,k_1} + \frac{1}{2}\delta_{*,k_2}$. On one hand if the decision-maker plays *a* then the state stays the same, and he receives a payoff of 1 if and only if the state is k_2 . On the other hand if he plays *b* then he receives a payoff of 0 but the probability to be in state k_2 increases.

We define the function r from $X \times A = \Delta(K) \times A$ to [0, 1] by r((p, 1-p), a) = 1-p and r((p, 1-p), b) = 0 for all $p \in [0, 1]$, and we define the transition \hat{q} from $X \times A$ to $\Delta_f(X)$ by :

$$\hat{q}((p, 1-p), a) = \delta_{(p,1-p)}$$
 and $\hat{q}((p, 1-p), b) = \delta_{(p/2,1-p/2)}$.

Then the standard MDP $\Psi = (\Delta(K), A, \hat{q}, r)$ is the MDP associated in the previous proof to Γ . This MDP is here deterministic, because the decision is in the dark.

The existence of a general uniform value is immediate here. Given $n \ge 1$, the strategy $\sigma = b^n a^{\infty}$ which plays n times b and then a for the rest of the game, guarantees a stage payoff of $(1-\frac{1}{2^n})$ from stage n+1 on, so the game has a general uniform value equal to 1. Finally if we consider the discounted evaluations, one can show that the speed of convergence of v_{λ} is here slower than λ :

$$v_{\lambda}(p_1) = 1 - \frac{\ln(\lambda)}{\ln(2)}\lambda + O(\lambda).$$

The partial observation allows for a speed of convergence slower than λ contrary to the perfect observation case where it is well known that the convergence is in $O(\lambda)$.

Remark 4.3. It is here unknown if the uniform value exists in pure strategies, i.e. if the behavior strategies $\sigma(p)$ of Theorem 4.1 can be chosen with values in A. This was already an open problem for the Cesàro-uniform value, that is when only evaluations of the form $\theta = \frac{1}{n} \sum_{t=1}^{n} \delta_t$ are considered (see Rosenberg *et al.* [36] and Renault [32] for different proofs requiring the use of behavioral strategies). In the present proof, there are two related places where the use of lotteries on actions is important. First in the proof of the convergence of the function $h_{T,n}$ (within the proof of Theorem 3.10), we used Sion's theorem in order to exchange a supremum and an infimum, and to do so the convexity of the set of strategies was required. Secondly when we prove that the extended transition is 1-Lipschitz (see Lemma 3.27), the coupling between the two distributions u and u' requires some randomization.

4.2 Zero-sum repeated games with an informed controller

We finally consider zero-sum repeated games with an informed controller. We start with a general model $\Gamma = (K, I, J, C, D, q, g)$ of zero-sum repeated game,

where we have 5 non empty finite sets: a set of states K, two sets of actions I and J and two sets of signals C and D, and we also have a transition mapping q from $K \times I \times J$ to $\Delta(K \times C \times D)$ and a payoff function g from $K \times I \times J$ to [0,1]. Given an initial probability π on $\Delta(K \times C \times D)$, the game $\Gamma(\pi) = \Gamma(K, I, J, C, D, q, g, \pi)$ is played as follows: at stage 1, a triple (k_1, c_1, d_1) is drawn according to π , player 1 learns c_1 and player 2 learns d_1 . Then simultaneously player 1 chooses an action i_1 in I and player 2 chooses an action j_1 in J. Player 1 gets a (unobserved) payoff $r(k_1, i_1, j_1)$ and player 2 the opposite payoff. Then a new triple (k_2, c_2, d_2) is drawn accordingly to $q(k_1, i_1, j_1)$. Player 1 observes c_2 , player 2 observes d_2 and the game proceeds to the next stage, etc...

A (behavioral) strategy for player 1 is a sequence $\sigma = (\sigma_t)_{t\geq 1}$ where for each $t \geq 1$, σ_t is a mapping from $(C \times I)^{t-1} \times C$ to $\Delta(I)$. Similarly a strategy for player 2 is a sequence of mappings $\tau = (\tau_t)_{t\geq 1}$ where for each $t \geq 1$, τ_t is a mapping from $(D \times J)^{t-1} \times D$ to $\Delta(J)$. We denote respectively by Σ and \mathcal{T} the set of strategies of player 1 and player 2. An initial distribution π and a couple of strategies (σ, τ) defines for each t a probability on the possible histories up to stage t, which can be uniquely extended to a probability on the set of infinite histories $(K \times C \times D \times I \times J)^{+\infty}$.

Given an evaluation θ , we define the θ -payoff of (σ, τ) in $\Gamma(\pi)$ as the expectation under $I\!\!P_{\pi,\sigma,\tau}$ of the payoff function,

$$\gamma_{\theta}(\pi, \sigma, \tau) = I\!\!E_{\pi, \sigma, \tau} \left(\sum_{t} \theta_t r(k_t, i_t, j_t) \right).$$

By Sion's theorem the game with θ -payoff has a value:

$$v_{\theta}(\pi) = \max_{\sigma \in \Sigma} \min_{\tau \in \mathcal{T}} \gamma_{\theta}(\pi, \sigma, \tau) = \min_{\tau \in \mathcal{T}} \max_{\sigma \in \Sigma} \gamma_{\theta}(\pi, \sigma, \tau),$$

and we can define the general limit value as in the MDP framework. Note that we do not ask the convergence to be uniform for all π in $\Delta(K \times C \times D)$, because we will later make some assumptions, in particular on the initial distribution.

Definition 4.4. The repeated game $\Gamma(\pi) = (K, I, J, C, D, q, g, \pi)$ has a general limit value $v^*(\pi)$ if $v_{\theta}(\pi)$ converges to $v^*(\pi)$ when $TV(\theta)$ goes to zero, i.e.:

$$\forall \varepsilon > 0, \exists \alpha > 0, \forall \theta, \ (TV(\theta) \le \alpha \implies (|v_{\theta}(\pi) - v^*(\pi)| \le \varepsilon)).$$

Definition 4.5. The repeated game $\Gamma(\pi)$ has a general uniform value if it has a general limit value $v^*(\pi)$ and for each $\varepsilon > 0$ one can find $\alpha > 0$ and a couple of strategies σ^* and τ^* such that for all evaluations θ with $TV(\theta) \leq \alpha$:

$$\forall \tau \in \mathcal{T}, \gamma_{\theta}(\pi, \sigma^*, \tau) \ge v^*(\pi) - \varepsilon \text{ and } \forall \sigma \in \Sigma, \gamma_{\theta}(\pi, \sigma, \tau^*) \le v^*(\pi) + \varepsilon.$$

Without further assumption, the general values may fail to exist. We will focus here on the case of a repeated game with an informed controller, as introduced in Renault [33]. The first assumption concerns the information of the first player. We assume that he can always reconstruct the current state and the signal of the second player from his own signal:

Assumption 4.6. There exist two mappings $\tilde{k} : C \to K$ and $\tilde{d} : C \to D$ such that, if E denotes $\{(k, c, d) \in K \times C \times D, \tilde{k}(c) = k, \tilde{d}(c) = d\}$, we have: $\forall (k, i, j) \in K \times I \times J, q(k, i, j)(E) = 1, and \pi(E) = 1.$

Moreover we will assume that only player 1 has a meaningful influence on the transitions, in the following sense.

Assumption 4.7. The marginal of the transition on $K \times D$ is not influenced by player 2's action. For k in K, i in I and j in J, we denote by $\bar{q}(k,i)$ the marginal of q(k,i,j) on $K \times D$.

The second player may influence the signal of the first player but he can not prevent him neither to learn the state nor to learn his own signal. Moreover he can not influence his own information, thus he has no influence on his beliefs about the state or about the beliefs of player 1 about his beliefs. A repeated game satisfying assumptions 4.6 and 4.7 is called a repeated game with an informed controller. It was proved in Renault [33] that for such games the Cesàro-uniform value (that is, when only evaluations of the form $\theta = \frac{1}{n} \sum_{t=1}^{n} \delta_t$ are considered) exists and we will extend it here to the general uniform value.

Example 4.8. We consider the simplest case of zero-sum repeated game with incomplete information introduced by Aumann and Maschler in the sixties (see reference [5]). It is defined by a finite family $(G^k)_{k \in K}$ of payoff matrices in $[0, 1]^{I \times J}$ and $p \in \Delta(K)$ an initial probability. At the first stage, some state k is selected according to p and told to player 1 only. The second player knows the initial distribution p but not the realization of the state. Then the matrix game G^k is repeated over and over. At each stage the players observe past actions but not their payoff (notice that player 1 can always reconstruct the payoff from the actions and the state). Formally it is a zero-sum repeated game $\Gamma = (K, I, J, C, D, q, g)$ as defined previously, with $C = K \times I \times J$ and $D = I \times J$, and for all $(k, i, j) \in K \times I \times J$, $g(k, i, j) = G^k(i, j)$ and $q(k, i, j) = \delta_{k,(k,i,j),(i,j)}$. For all $p \in \Delta(K)$, we denote by $\Gamma(p)$ the game where the initial probability $\pi \in \Delta(K \times C \times D)$ is given by $\pi = \sum_{k \in K} p(k) \delta_{k,(k,i_0,j_0),(i_0,j_0)}$ with $(i_0, j_0) \in I \times J$ fixed. For each n, we denote by $v_n(p)$ the value of the n-stage game with initial

For each n, we denote by $v_n(p)$ the value of the n-stage game with initial probability p, where the payoff is the expected average of the first n payoffs. The value satisfies the standard recursive formula:

$$v_n(p) = \sup_{a \in \Delta(I)^K} \left(\frac{1}{n} r(p, a) + \frac{n-1}{n} \sum_{i \in I} a(p)(i) v_{n-1}(\hat{q}(p, a, i)) \right),$$

where $a^k \in \Delta(I)$ represents the lottery on actions played by player 1 if the state is $k, a(p)(i) = \sum_{k \in K} p^k a^k(i)$ is the probability that player 1 plays i, r(p, a) = $\min_j(\sum_k p^k G^k(a^k, j))$ is the minimal expected payoff for player 1, and $\hat{q}(p, a, i)$ is the conditional belief on $\Delta(K)$ given p, a, i:

$$\hat{q}(p, a, i) = \left(\frac{p(k)a^k(i)}{a(p)(i)}\right)_k$$

Starting from a belief p about the state, if player 2 observes action i and knows that the distribution of actions of player 1 is a, then he updates his beliefs to $\hat{q}(p, a, i)$. Aumann and Maschler have proved that the limit value exists and is characterized by

$$v^* = \operatorname{cav} f^* = \inf\{v : \Delta(K) \to [0, 1], v \text{ concave } v \ge f^*\},\$$

where $f^*(p) = Val\left(\sum_k p^k G^k\right)$ for all $p \in \Delta(K)$. The function f^* is the value of the game, called the non-revealing game, where player 1 is forbidden to use his information.

Theorem 4.9. A zero-sum repeated game with an informed controller has a general uniform value.

Proof of Theorem 4.9: Assume that $\Gamma(\pi) = (K, I, J, C, D, q, g, \pi)$ is a repeated game with an informed controller, i.e. that assumptions 4.6 and 4.7 are satisfied. The proof will consist of 5 steps. First we introduce an auxiliary standard Markov Decision Process $\Psi(\hat{\pi})$ on the state space $X = \Delta(K)$. Then we show that for all evaluations θ , the repeated game $\Gamma(\pi)$ and the MDP $\Psi(\hat{\pi})$ have the same θ -value. In step 3 we check that the MDP satisfies the assumption of Theorem 3.20 so it has a general limit value and a general uniform value v^* . As a consequence the repeated game has a general limit value $v^*(\pi)$. Then we prove that player 1 can use an ϵ -optimal strategy of the auxilliary MDP in order to guarantee $v^*(\pi) - \epsilon$ in the original game. Finally we prove that Player 2 can play by blocks in the repeated game in order to guarantee $v^*(\pi) + \epsilon$. And we obtain that $v^*(\pi)$ can be guaranteed by both players in the repeated game, so it is the general uniform value of $\Gamma(\pi)$.

For every $P \in \Delta(K \times C \times D)$, we denote by \overline{P} the marginal of P on $K \times D$ and we put $\hat{P} = \psi_D(\overline{P})$ where ψ_D is the disintegration with respect to D (recall Theorem 2.22): for all $\mu \in \Delta(K \times D)$, $\psi_D(\mu) = \sum_{d \in D} \mu(d) \delta_{\mu(.|d)}$.

Step 1: We put $X = \Delta(K)$ and $A = \Delta(I)^K$ and for every p in X, a in A and

b in $\Delta(J)$, we define:

$$r(p, a, b) = \sum_{\substack{(k,i,j) \in K \times I \times J}} p^k a^k(i) b(j) g(k, i, j) \in [0, 1],$$

$$r(p, a) = \inf_{b \in \Delta(J)} r(p, a, b) = \inf_{j \in J} r(p, a, j),$$

$$\overline{q}(p, a) = \sum_{\substack{(k,i) \in K \times I}} p^k a^k(i) \overline{q}(k, i) \in \Delta(K \times D),$$

$$\hat{q}(p, a) = \psi_D(\overline{q}(p, a)) = \sum_{d \in D} \overline{q}(p, a)(d) \delta_{\hat{q}(p, a, d)} \in \Delta_f(X).$$

Here $\hat{q}(p, a, d) \in \Delta(K)$ is the belief of the second player on the new state after observing the signal d and knowing that player 1 has played a at p:

$$\forall k' \in K, \hat{q}(p, a, d)(k') = \frac{\overline{q}(p, a)(k', d)}{\overline{q}(p, a)(d)} = \frac{\sum_{k} p^{k} q(k, a(k))(k', d)}{\sum_{k} p^{k} q(k, a(k))(d)}.$$

We define the auxiliary MDP $\Psi = (X, A, \hat{q}, r)$, and denote the θ -value in the MDP by \hat{v}_{θ} . The MDP with initial state $\hat{\pi}$ has strong links with the repeated game $\Gamma(\pi)$.

Step 2: By proposition 4.23, part b) in Renault [33], we have for all evaluations θ with finite support:

$$v_{\theta}(\pi) = \hat{v}_{\theta}(\hat{\pi}).$$

The proof relies on the same recursive formula satisfied by v and \hat{v} , and the equality can be easily extended to any evaluation θ .

$$\forall \theta \in \Delta(\mathbb{N}^*), \forall p \in X, \ v_{\theta}(p) = \sup_{a \in A} \inf_{b \in B} \left(\ \theta_1 r(p, a, b) + (1 - \theta_1) v_{\theta^+}(\hat{q}(p, a)) \right).$$

where v_{θ^+} is naturally linearly extended to $\Delta_f(X)$. As a consequence if $\Psi(\hat{\pi})$ has a general limit value so does the repeated game $\Gamma(\pi)$.

Step 3: Let us check that Ψ satisfies the assumption of Theorem 3.20. Consider p, p' in X, a in A, and $\alpha \ge 0$ and $\beta \ge 0$. We have:

$$\begin{aligned} |\alpha r(p,a) - \beta r(p',a)| &\leq \sup_{b \in \Delta(J)} |\alpha r(p,a,b) - \beta r(p',a,b)| \\ &\leq \sup_{b \in \Delta(J)} |\sum_{k \in K} \alpha p^k g(k,a^k,b) - \beta p'^k g(k,a^k,b)| \\ &\leq \sup_{b \in \Delta(J)} \sum_{k \in K} |\alpha p^k - \beta p'^k| = \|\alpha p - \beta p'\|_1. \end{aligned}$$

Moreover, let $\varphi : \Delta(K) \longrightarrow \mathbb{R}$ be in D_1 .

$$\begin{aligned} |\alpha\varphi(\hat{q}(p,a)) - \beta\varphi(\hat{q}(p',a))| &= \sum_{d\in D} \left(\alpha\bar{q}(p,a)(d)\varphi(\hat{q}(p,a,d)) - \beta\bar{q}(p',a)(d)\varphi(\hat{q}(p',a,d))\right) \\ &\leq \sum_{d\in D} \|\alpha\;\bar{q}(p,a)(d)\;\hat{q}(p,a,d) - \beta\;\bar{q}(p',a)(d)\;\hat{q}(p',a,d)\|_{1} \\ &\leq \sum_{d\in D} \|\alpha\;(\bar{q}(p,a)(k',d))_{k'} - \beta\;(\bar{q}(p',a)(k',d))_{k'}\|_{1} \\ &\leq \sum_{d\in D} \sum_{k\in K} \|\alpha p^{k}\;(\bar{q}(k,a)(k',d))_{k'} - \beta p'^{k}\;(\bar{q}(k,a)(k',d))_{k'}\|_{1} \\ &\leq \sum_{d\in D} \sum_{k'\in K} \sum_{k\in K} \bar{q}(k,a)(k',d)|\alpha p^{k} - \beta p'^{k}| = \|\alpha p - \beta p'\|_{1}. \end{aligned}$$

So $\Psi = (X, A, \hat{q}, r)$ has a general limit value and a general uniform value that we denote by v^* . As a consequence, $\Gamma(\pi)$ has a general limit value $v^*(\pi)$.

Step 4: Given $\varepsilon > 0$, there exists $\alpha > 0$ and a strategy σ in the MDP $\Psi(\hat{\pi})$ such that the θ -payoff in the MDP is large: $\hat{\gamma}_{\theta}(\hat{\pi}, \sigma) \geq v^*(\pi) - \varepsilon$ whenever $TV(\theta) \leq \alpha$. Moreover if we look at the end of the proof of Theorem 3.20 we can choose σ to be induced by a deterministic play in the Gambling House $\hat{\Gamma}$ with state space $Z = \Delta_f(X) \times [0, 1]$. As a consequence one can mimic σ to construct a strategy σ^* in the original repeated game $\Gamma(\pi)$ such that: $\forall \tau \in \mathcal{T}, \gamma_{\theta}(\pi, \sigma^*, \tau) \geq v^*(\pi) - \varepsilon$ whenever $TV(\theta) \leq \alpha$.

Step 5: Finally we show that player 2 can also guarantee the value v^* in the repeated game Γ . Note that in the repeated game he can not compute the state variable in $\Delta(K)$ without knowing the strategy of player 1. Nevertheless he has no influence on the transition function so playing independently by large blocks will be sufficient for him in order to guarantee $v^*(\pi)$. We use the following characterization of the value proved in Renault [33]:

$$v^*(\pi) = \inf_n \sup_m v_{m,n}(\pi).$$

where $v_{m,n}$ is the value of the game with payoff function the Cesàro mean of the stage payoffs between stages m + 1 and m + n. We proceed as in proposition 4.22 of Renault [33]. Fix $n_0 \geq 1$, then we consider the strategy τ^* which for each $j \in \mathbb{N}$, plays optimally in the game with evaluation the Cesàro mean of the payoffs on the block of stages $B^j = \{n_0(j-1) + 1, ..., n_0 j\}$. Since player 2 does not influence the state, τ^* is well defined and guarantees $\sup_{t\geq 0} v_{t,n_0}(z)$ on each block B^j .

Let θ be an evaluation and σ be a strategy of player 1. For each $j \geq 1$, denote

by θ_j the minimum of θ on the block B^j . We have

$$\gamma_{\theta}(\pi, \sigma, \tau^{*}) = \sum_{j=1}^{+\infty} I\!\!E_{\pi, \sigma, \tau^{*}} \left(\sum_{t=(j-1)n_{0}+1}^{jn_{0}} \theta_{t} g(k_{t}, a_{t}, b_{t}) \right)$$

$$\leq \sum_{j=1}^{+\infty} n_{0} \underline{\theta}_{j} \sup_{t \ge 0} v_{t,n_{0}}(\pi) + n_{0} \sum_{t=1}^{+\infty} |\theta_{t+1} - \theta_{t}|$$

$$\leq \sup_{t \ge 0} v_{t,n_{0}}(\pi) + n_{0} TV(\theta).$$

Given ϵ , there exists n_0 such that $\sup_{t\geq 0} v_{t,n_0}(\pi) \leq v^*(\pi) + \epsilon$. Fix $\alpha = \frac{\epsilon}{n_0}$ and τ^* defined as before then for all θ such that $TV(\theta) \leq \alpha$, we have

$$\sup_{\sigma\in\Sigma}\gamma_{\theta}(\pi,\sigma,\tau^*)\leq v^*(\pi)+2\epsilon,$$

and this concludes the proof of Theorem 4.9.

Example 4.10. The computation of the value is in general a difficult problem, as shown by the next example introduced in Renault [31] and studied by Hörner *et al.* [21]. In this example the value exists but has been computed only for some values of the parameter. The set of states is $K = \{k_1, k_2\}$, the set of actions of player 1 is $I = \{T, B\}$, the set of actions of player 2 is $J = \{L, R\}$, and the payoff of player 1 is given by:

$$\begin{array}{cccccc}
L & R & & L & R \\
T & \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix} & \text{and} & T & \begin{pmatrix} 0 & 0 \\ 0 & 1 \end{pmatrix} \\
k_1 & & k_2
\end{array}$$

The sequence of states follows an exogeneous Markov chain, with initial probability (1/2, 1/2) and transition matrix $\begin{pmatrix} p & (1-p) \\ (1-p) & p \end{pmatrix}$, where p is a given parameter. At the beginning of every stage, only player 1 observes the current state in K, and at the end of each stage the actions played are observed. (with the previous notations $C = K \times I \times J$, $D = I \times J$, and $q(k, i, j) = p \, \delta_{k,(k,i,j),(i,j)} + (1-p) \, \delta_{k',(k',i,j),(i,j)}$ for all k in $K, k' \in K \setminus \{k\}, i$ in I and j in J).

For each value of the parameter $p \in [0, 1]$, we have a repeated game Γ^p and by symmetry it is sufficient to study the case $p \in [1/2, 1]$. If p = 1 we are in the setup of Example 4.8 and the value is clearly 1/4. Hörner *et al.* [21] proved that for $p \in [1/2, 2/3)$, the value is $v_p = \frac{p}{4p-1}$. For $p \ge 2/3$ the value is not known, except for p^* , the solution of $9p^3 - 12p^2 + 6p - 1 = 0$, where one has $v_{p^*} = \frac{p^*}{1-3p^*+6(p^*)^2}$.

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References

- Arapostathis A., Borkar V., Fernández-Gaucherand E., Ghosh M., Marcus S.: Discrete-time controlled Markov processes with average cost criterion: a survey. SIAM J. Control Opt. 31(2), 282–344 (1993)
- [2] Ash R.: Real analysis and probability, vol. 239. Academic Press New York (1972)
- [3] Aström K.: Optimal control of Markov processes with incomplete state information. J. Math. Anal. Appl. 10, 174–205 (1965)
- [4] Aubin J.: Applied abstract analysis. John Wiley & Sons (1977)
- [5] Aumann R., Maschler M., Stearns R.: Repeated games with incomplete information. The MIT press (1995)
- [6] Bellman R.: A Markovian decision problem. J. Math. Mech. (1957)
- Bewley T., Kohlberg E.: The asymptotic theory of stochastic games. Math. Oper. Res. 1(3), 197–208 (1976)
- [8] Birkhoff G.: Proof of the ergodic theorem. Proc. Natl. Acad. Sci. USA 17(12), 656-660 (1931)
- [9] Blackwell D.: Discrete dynamic programming. Ann. Math. Statist. 33, 719–726 (1962)
- [10] Borkar V.: Average cost dynamic programming equations for controlled Markov chains with partial observations. SIAM J. Control Opt. 39(3), 673-681 (2000)
- [11] Borkar V.: Convex analytic methods in Markov decision processes. In: Feinberg E., Shwartz A.(eds) Handbook of Markov decision processes, 347–375. Kluwer (2002)
- [12] Borkar V.: Dynamic programming for ergodic control of Markov chains under partial observations: a correction. SIAM J. Control Opt. 45(6), 2299–2304 (2007)
- [13] Choquet G.: Existence et unicité des représentations intégrales au moyen des points extrémaux dans les cônes convexes. Sém. Bourbaki 4, 33–47 (1956)
- [14] Denardo E., Fox B.: Multichain Markov renewal programs. SIAM J. Appl. Math. 16(3), 468–487 (1968)
- [15] Dubins L., Savage L.: How to gamble if you must: Inequalities for stochastic processes. McGraw-Hill New York (1965)
- [16] Dudley R. M.: Real analysis and probability, vol. 74. Cambridge University Press (2002)

- [17] Filar J., Sznajder R.: Some comments on a theorem of Hardy and Littlewood. J. Optim. Theory Appl. 75(1), 201–208 (1992)
- [18] Harsanyi J.: Games with incomplete information played by "Bayesian" players, I-III. part I. The basic model. Manag. Sci. 14(3), 159–182 (1967)
- [19] Hernández-Lerma O.: Adaptive Markov control processes, vol. 79. Springer New York (1989)
- [20] Hordijk A., Kallenberg L.: Linear programming and Markov decision chains. Manag.Sci. 25(4), 352–362 (1979)
- [21] Hörner J.,Rosenberg D.,Solan E.,Vieille N.: On a Markov game with one-sided information. Oper. Res. 58(4), 1107–1115 (2010)
- [22] Kallenberg L.: Survey of linear programming for standard and nonstandard Markovian control problems. part i: Theory. Math. Methods Oper. Res. 40(1), 1–42 (1994)
- [23] Lehrer E., Sorin S.: A uniform Tauberian theorem in dynamic programming, Math. Methods Oper. Res. 17(2), 303–307 (1992)
- [24] Maitra A., Sudderth W.: Discrete gambling and stochastic games, vol. 32. Springer Verlag (1996)
- [25] McShane E.J.: Extension of range of functions. Bull. Amer. Math. Soc. (N.S.) 40.12, 837–842 (1934)
- [26] Mertens J.-F., Neyman A.: Stochastic games. Internat. J. Game Theory 10(2), 53–66 (1981)
- [27] Mertens J.-F., Zamir S.: Formulation of Bayesian analysis for games with incomplete information, International J. Game Theory 14(1), 1–29 (1985)
- [28] Mertens J.-F.: Repeated games, in Proceedings of the International Congress of Mathematicians, Vol. 1, 2 (Berkeley, Calif., 1986), (Providence, RI), pp. 1528– 1577, Amer. Math. Soc. (1987)
- [29] Mertens J.F., Sorin S., Zamir S.: Repeated games, CORE Discussion Papers, vol. 9420,9421,9422 (1994)
- [30] Neyman A.: Existence of optimal strategies in markov games with incomplete information. Internat. J. Game Theory 37(4), 581–596 (2008)
- [31] Renault J.: The value of Markov chain games with lack of information on one side. Math. Oper. Res. 31(3), 490–512 (2006)
- [32] Renault J.: Uniform value in dynamic programming. J. Eur. Math. Soc. 13(2), 309–330 (2011)
- [33] Renault J.: The value of repeated games with an informed controller. Math. Oper. Res. 37(1), 154–179 (2012)

- [34] Renault J.: General long-term values in dynamic programming. To appear in J. Dynam. Games (2013)
- [35] Rhenius D.: Incomplete information in Markovian decision models. Ann. Statist. 2(6), 1327–1334 (1974)
- [36] Rosenberg D., Solan E., Vieille N.: Blackwell optimality in Markov decision processes with partial observation. Ann. Statist. 30(4), 1178–1193 (2002)
- [37] Rosenberg D., Solan E., Vieille N.: Stochastic games with a single controller and incomplete information. SIAM J. Control Optim. 43(1), 86–110 (electronic) (2004)
- [38] Runggaldier W. J., Stettner L.: On the construction of nearly optimal strategies for a general problem of control of partially observed diffusions, Stochastics: Intern. J.Prob. Stochastic Proc. 37(1-2), 15–47 (1991)
- [39] Sawaragi Y., Yoshikawa T.: Discrete-time Markovian decision processes with incomplete state observation. Ann. Math. Statist. 41(1), 78–86 (1970)
- [40] Shapley L.: Stochastic games. Proc. Natl. Acad. Sci. U. S. A. 39(10), 1095–1100 (1953)
- [41] Sorin S.: A first course on zero-sum repeated games, vol. 37. Springer (2002)
- [42] Von Neumann J.: Proof of the quasi-ergodic hypothesis. Proc. Nat. Acad. Sci. U.S.A. 18(1), 70–82 (1932)