

Matching random colored points with rectangles

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Abstract

Let $S \subset [0, 1]^2$ be a set of n points, randomly and uniformly selected. Let $R \cup B$ be a random partition, or coloring, of S in which each point of S is included in R uniformly at random with probability $1/2$. We study the random number $M(n)$ of points of S that are covered by the rectangles of a maximum strong matching of S with axis-aligned rectangles. The matching consists of closed rectangles that cover exactly two points of S of the same color. A matching is strong if all its rectangles are pairwise disjoint. We prove that almost surely $M(n) \geq 0.83n$ for n large enough. Our approach is based on modeling a deterministic greedy matching algorithm, that runs over the random point set, as a Markov chain.

1 Introduction

Given a point set $S \subset \mathbb{R}^2$ of n points, and a class \mathcal{C} of geometric objects, a *geometric matching* of S is a set $M \subseteq \mathcal{C}$ such that each element of M contains exactly two points of S and every point of S lies in at most one element of M . A geometric matching is *strong* if the geometric objects are pairwise disjoint, and *perfect* if every point of S belongs to (or is covered by) some element of M . This type of geometric matching problems was considered by Ábrego et al. [1], who studied the existence and properties of matchings for point sets in the plane when \mathcal{C} is the class of axis-aligned squares, or the class of disks.

Let $S = R \cup B \subset \mathbb{R}^2$ be a set of n colored points in the plane, each point colored red or blue, where

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R and B are the sets of red and blue points, respectively. A geometric matching of S is called *monochromatic* if all matching objects cover points of the same color, and *bichromatic* if all matching objects cover points of different colors. For example, monochromatic matchings of two-colored point sets in the plane with straight segments have been studied [4, 5]. In the case of bichromatic matchings with straight segments, a classical result in discrete geometry asserts that for any planar point set S consisting of n red points and n blue points in general position (i.e., no three points of S are collinear) there exists a perfect, strong bichromatic matching of S with straight segments.

In this paper, we consider strong monochromatic matchings with axis-aligned rectangles. Every rectangle will be axis-aligned and a closed set.

Caraballo et al. [2] studied monochromatic strong matchings of S with rectangles from the algorithmic point of view. That is, the problem of finding a monochromatic strong matching of S with the maximum number of rectangles; proving that the problem is NP-hard and giving a polynomial-time 4-approximation algorithm. As noted by Caraballo et al., this problem is a special case of the Maximum Independent Set of Rectangles problem (MISR): Given a finite set \mathcal{R} of rectangles in the plane, find a subset $\mathcal{R}' \subseteq \mathcal{R}$ of maximum cardinality, denoted $\alpha(\mathcal{R})$, such that every pair of rectangles in \mathcal{R}' are disjoint.

Indeed, suppose that we want to find a monochromatic matching of S with the maximum number of rectangles. For every distinct $p, q \in \mathbb{R}^2$, let $D(p, q)$ be the minimum axis-aligned rectangle that encloses p and q . Let $\mathcal{R}(S)$ be the set of all rectangles $D(p, q)$ such that $p, q \in S$, p and q have the same color, and $D(p, q)$ contains no points of S different from p and q . Finding a monochromatic strong matching of S with the maximum number of rectangles is equivalent to finding in $\mathcal{R}(S)$ a maximum subset of pairwise disjoint rectangles, whose size is $\alpha(\mathcal{R}(S))$. That is, to solving the MISR problem in $\mathcal{R}(S)$.

We study monochromatic strong matchings of S with rectangles from the combinatorial point of view, and from this point forward, every rectangle will cover precisely two points of S . Point sets $S = R \cup B$ exist

in which no matching rectangle is possible (e.g., S is a color-alternating sequence of points on the line $y = x$), and point sets in which a perfect strong matching with rectangles exists (e.g., an even number of red points in the negative part of the line $y = x$, and an even number of blue points in the positive part). These two extreme cases show that it is not worth studying the number $\alpha(\mathcal{R}(S))$ for fixed, or given, colored point sets S . Instead, we want to study $\alpha(\mathcal{R}(S))$ when S is a random point set in the square $[0, 1]^2$, in which the positions of the n points of S are random and the color of each point of S is also random. Formally:

Let $n > 0$, and let $S \subset [0, 1]^2$ be a set of n points, randomly and uniformly selected. Let $R \cup B$ be a random partition (i.e., coloring) of S in which each point of S is included in R uniformly at random with probability $1/2$. We study the random variable $M(n) = 2 \cdot \alpha(\mathcal{R}(S))$ equal to the number of points of S that are covered by the rectangles of a maximum monochromatic strong matching of S with rectangles.

Given a set S of n points, randomly and uniformly selected in the square $[0, 1]^2$, Chen et al. [3] studied a similar variable: the random variable $\alpha(D(S))$, where $D(S)$ is the random graph with vertex set S and two points $p, q \in S$ define an edge if and only if $D(p, q) \cap S = \{p, q\}$. Here, $\alpha(D(S))$ denotes the size of a maximum independent set of $D(S)$.

One result of Chen et al. [3, Theorem 1] states that if n tends to infinity, then we have $\alpha(D(S)) = O(n(\log^2 \log n) / \log n)$ with probability tending to 1. This result implies that if $C(n)$ denotes the number of points of S that are covered by a maximum monochromatic matching of S with rectangles, where the rectangles may overlap (i.e., the matching is not necessarily strong), then $C(n) = n - o(n)$ with probability tending to 1. In fact, let M' be a maximum monochromatic matching of S with rectangles, where M' is not necessarily strong, and let $S' \subset S$ be the points not covered by M' . Note that at least $|S'|/2$ points of S' have the same color, and they form an independent set in the graph $D(S)$. Then, with probability tending to 1, we have that M' covers at least $n - |S'| = n - O(n(\log^2 \log n) / \log n) = n - o(n)$ points.

2 Preliminaries

Since for matching S with rectangles, only the left-to-right and bottom-to-top orders of S are relevant, and since the probability that two points of S are in the same vertical or horizontal line is zero, we consider S equal to the point set $S_\pi = \{(i, \pi(i)) \mid i = 1, 2, \dots, n\}$, where $\pi : \{1, 2, \dots, n\} \rightarrow \{1, 2, \dots, n\}$ is a randomly and uniformly selected permutation. This assumption was also done by Chen et al. [3].

We have implemented a program that, given n , generates a uniform random permutation π , and selects the color of each $p \in S_\pi$ (red or blue) randomly and

k	$n = 1000$		$n = 10000$	
	mean	sdev	mean	sdev
1	0.6653	0.0175	0.6673	0.0052
2	0.7948	0.0104	0.7934	0.0036
3	0.8301	0.0097	0.8304	0.0034
4	0.8555	0.0094	0.8562	0.0028
5	0.8727	0.0090	0.8736	0.0026
6	0.8860	0.0087	0.8864	0.0026
∞	0.9724	0.0062	0.9780	0.0022

Table 1: The experimental results obtained when running the greedy matching algorithm for $n \in \{1000, 10000\}$, parameterized with $k \in [1..6]$, or not parameterized ($k = \infty$). For each combination n, k , we run the algorithm 100 times, and measured the mean and standard deviation of the ratio between the total number of matched points and n .

uniformly. The program then runs a deterministic algorithm on $S_\pi = R \cup B$ that greedily finds a maximum independent subset of rectangles in $\mathcal{R}(S_\pi)$. The algorithm iterates the points of S_π from left to right, and for each point p in the iteration, it performs the following action: If p is not matched with any point prior to p in the iteration, it finds the first point q to the right of p such that $D(p, q) \in \mathcal{R}(S_\pi)$ and $D(p, q)$ has empty intersection with all matching rectangles already reported. If q exists, the algorithm reports $D(p, q)$ as a matching rectangle. In any case, regardless of whether q exists, the algorithm continues the iteration to the next unmatched point p .

For large n , say $n = 10000$, the implemented algorithm reports a matching covering approximately $\frac{97}{100}n$ of the points. Then, it seems that $M(n) \geq \frac{97}{100}n$ for n large enough and probability close to 1. Analyzing the algorithm, when run over the random S_π , seems to be a good approach for obtaining a high lower bound for $M(n)$. One way to analyze the algorithm is to consider a parameterized version of it, with a parameter k , such that each unmatched point p finds its match point q among only the next k points of S_π to the right of p . Let \mathcal{A}_k denote this parameterized algorithm. For experimental results, see Table 1.

We show how to model (an adaptation of) \mathcal{A}_k as a Markov chain, for any $k \in \{1, 2, \dots\}$. Then, we show that \mathcal{A}_3 almost surely guarantees $M(n) \geq \frac{83}{100}n$, for n large enough, by computing the stationary distribution of the Markov chain and applying the Ergodic theorem. See [6] for the theory on Markov chains.

3 The Markov chains

We consider $S = S_\pi$, and whenever we say point i , for $i \in \{1, 2, \dots, n\}$, or just i when it is clear from the context, we are referring to the point $p_i := (i, \pi(i)) \in S$. Let $\text{color}(i) \in \{R, B\}$ denote the color of point i .

Let $k \in \{1, 2, 3, \dots\}$ be a constant, and let $\tilde{\mathcal{A}}_k$ be the following adaptation of algorithm \mathcal{A}_k , consisting

in the next idea: Suppose that \mathcal{A}_k matches points i and j , with $i < j \leq i + k$, when the iteration of S_π is on point i . $\tilde{\mathcal{A}}_k$ iterates S_π from left to right, and will also match i and j but, in contrast with \mathcal{A}_k , when the iteration is on j , or on a point to the right of j . Using $\tilde{\mathcal{A}}_k$ instead of \mathcal{A}_k , allows us to describe in a more compact way the states of the memory of the algorithm during the iteration of the elements of S_π .

Let $E(j)$ be the data structure associated with point $j \in \{1, 2, \dots, n\}$, that is maintained by $\tilde{\mathcal{A}}_k$ during the iteration of S_π . For any j , let $i = i(j)$ be the smallest element in the set $\{\max(1, j - (k - 1)), \dots, j\}$ such that the point i is not matched, and each point in $\{i + 1, \dots, j\}$ is matched with a point to the left of i or is not yet matched. If i exists, then $E(j)$ consists of the following elements:

- The set $U(j) \subseteq \{i, i + 1, \dots, j\}$ of the points that are not matched, with $i \in U(j)$.
- The set $\text{Rect}(j)$ of the (pairwise disjoint) rectangles that match the points in $\{i + 1, \dots, j\} \setminus U(j)$ with points to the left of i .
- The number $f(j)$ of points of S_π that are matched while the iteration is at point j .

If i does not exist, then $E(j)$ consists of the same three above elements with $U(j) = \emptyset$ and $\text{Rect}(j) = \emptyset$.

For $j = 1$, we have $U(1) = \{1\}$, $\text{Rect}(1) = \emptyset$, and $f(1) = 0$. We show now how to obtain $E(j + 1)$ from $E(j)$, for any $j \in \{1, \dots, n - 1\}$. First, we match points i and $j + 1$ if and only if $j + 1 \leq i + k$, $\text{color}(i) = \text{color}(j + 1)$, and the rectangle $D(p_i, p_{j+1})$ does not overlap any rectangle in $\text{Rect}(j)$. After that, we match other points in $(U(j) \setminus \{i\}) \cup \{j + 1\}$ if and only if i was matched in the previous step, or we have finished with point i . We say that we have *finished* with point i if there do not exist more chances for point i to be matched, which is equivalent to $i + k \leq j + 1$. This final matching procedure consists in running the original algorithm \mathcal{A}_k with input the points $\{i + 1, \dots, j, j + 1\}$, but with the extra condition that the algorithm terminates if the current point t on the iteration of $\{i + 1, \dots, j, j + 1\}$ from left to right, cannot be matched with any other one to its right. This is because t must find its match among the points in $\{j + 2, \dots, t + k\}$, before any matching between points in $\{t + 1, \dots, j + 1\}$ occurs. We set $f(j + 1)$ equal to the total number of points matched in the above steps. Obtaining $U(j + 1)$ and $\text{Rect}(j + 1)$ is straightforward.

Let $j \in \{1, 2, \dots, n\}$. The *state* of $E(j)$ is the 2-tuple formed by: As first component, (a certificate of) the relative positions between the points of $U(j)$ and the rectangles of $\text{Rect}(j)$, together with the color of each point of $U(j)$. If the leftmost point is blue, then we switch the color of every point such that the leftmost one is always red. As second component, $f(j)$. We say that two states e and e' are *equal* (i.e., $e = e'$)

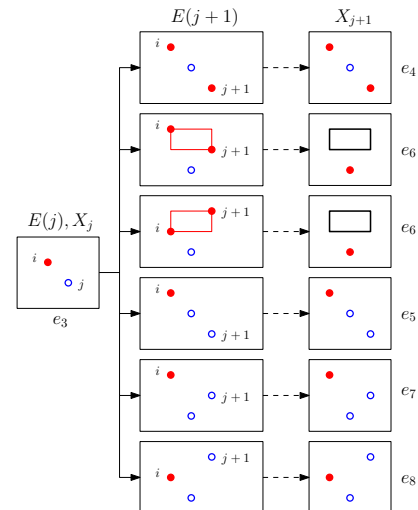


Figure 1: Example of the data structure $E(j)$, its state $X_j = e_3$, and the states in the neighborhood $N(e_3) = \{e_4, e_5, e_6, e_6, e_7, e_8\}$ corresponding to $E(j + 1)$, for each position and color of point $j + 1$. Note that $f(e_6) = 2$, and $f(e) = 0$ for all $e \in \{e_4, e_5, e_7, e_8\}$.

if: (i) the first components are equal, or one first component is symmetric to the other in the vertical direction, and (ii) the second components are equal.

Let $\mathcal{E} = \{e_1, e_2, \dots, e_N\}$ be the set of all possible states of $E(j)$, which is a finite set, and let $X_j \in \mathcal{E}$ be the random variable equal to the state of $E(j)$. Let $e \in \mathcal{E}$ be a state, and assume that e is the state of $E(j)$ for some j . Let $f(e) = f(j)$ (with abuse of notation), and let $N(e)$ be the *neighborhood* of e , which is the multiset consisting of the state of $E(j + 1)$ for every color and every different relative position, with respect to the elements of both $U(j)$ and $\text{Rect}(j)$, of point $j + 1$. See for example Figure 1.

Lemma 1 Let $e, e' \in \mathcal{E}$ be two states. For every $j \geq 2$, we have:

$$\text{Prob}(X_{j+1} = e' \mid X_j = e) = \frac{m}{2(|U(j)| + 2|\text{Rect}(j)| + 1)},$$

where m is the multiplicity of e' in $N(e)$.

Proof. Through each point of $U(j)$ draw a horizontal line, and for each rectangle of $\text{Rect}(j)$ draw a horizontal line through the top side and a horizontal line through the bottom side. Each of these $K = |U(j)| + 2|\text{Rect}(j)|$ lines goes through a different element of S_π , subdividing the plane into $K + 1$ strips. Since the point $j + 1$ is to the right of every point of $U(j)$ and every rectangle of $\text{Rect}(j)$, its relative position w.r.t. the elements of $U(j)$ and $\text{Rect}(j)$ is to be in one of these strips, and this happens with probability $1/(K + 1)$. Furthermore, the color of point $j + 1$ is given with probability $1/2$. The lemma follows. \square

Note that $\text{Prob}(X_{j+1} = x_{j+1} \mid X_j = x_j, \dots, X_1 = x_1) = \text{Prob}(X_{j+1} = x_{j+1} \mid X_j = x_j)$ for all

$x_1, \dots, x_{j+1} \in \mathcal{E}$ such that $\text{Prob}(X_j = x_j, \dots, X_1 = x_1) > 0$. Thus, $(X_n)_{n \geq 1}$ is a Markov chain, denoted \mathcal{C}_k , over the set $\mathcal{E} = \{e_1, e_2, \dots, e_N\}$ of states. Let P be the transition matrix, of dimensions $N \times N$, such that $P_{i,j} = \text{Prob}(X_{\ell+1} = e_j \mid X_\ell = e_i)$. The key observation is that the total number of points matched by $\tilde{\mathcal{A}}_k$ is precisely $M_k(n) := \sum_{j=1}^n f(X_j)$.

A Markov chain is *irreducible* if with positive probability any state can be reached from any other state [6]. It can be proved that \mathcal{C}_k is irreducible. Since \mathcal{C}_k has a finite state set, it has a unique stationary distribution $s = (s_1, s_2, \dots, s_N)$, which is the solution of the system $s = s \cdot P$, $s_1 + s_2 + \dots + s_N = 1$ of linear equations [6]. Furthermore, since $f(e) \in \{0, 2, 4, \dots, 2\lceil \frac{k+1}{2} \rceil\}$ for all $e \in \mathcal{E}$, the function f is bounded and then the Ergodic theorem ensures

$$\lim_{n \rightarrow \infty} \frac{M_k(n)}{n} = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n f(X_j) = \sum_{i=1}^N s_i f(e_i),$$

almost surely [6]. Let $\alpha_k = \sum_{i=1}^N s_i f(e_i)$. We then arrive to the main result of this paper:

Theorem 2 *Let $\pi : \{1, 2, \dots, n\} \rightarrow \{1, 2, \dots, n\}$ be a uniform random permutation. Let $S_\pi = \{(i, \pi(i)) \mid i = 1, 2, \dots, n\}$ be a random point set, where the color (red or blue) of each point of S_π is selected randomly and uniformly with probability 1/2. Let $k \in \{1, 2, 3, \dots\}$ be a constant. For all constant $\varepsilon > 0$ and n large enough, almost surely the number $M_k(n)$ of points of S_π that are matched by the algorithm $\tilde{\mathcal{A}}_k$ satisfies $M_k(n) \geq (\alpha_k - \varepsilon)n$.*

4 The Markov chain for $k = 3$

Using algorithm $\tilde{\mathcal{A}}_3$, we give a precise value for α_3 . In Table 2, we describe the states, and the transitions between the states, of the Markov chain \mathcal{C}_3 . Since $f(e) = 2$ for all $e \in \{e_2, e_6, e_9, e_{10}, e_{16}, e_{17}, e_{18}\}$, $f(e_{11}) = 4$, $f(e) = 0$ for all other state e , and the stationary distribution $s = (s_1, \dots, s_{18})$ satisfies

$$s_2 = \frac{167959}{816233}, s_6 = \frac{69640}{816233}, s_9 = \frac{6800}{816233}, s_{10} = \frac{58650}{816233},$$

$$s_{11} = \frac{13600}{816233}, s_{16} = \frac{5950}{816233}, s_{17} = \frac{1360}{816233}, s_{18} = \frac{1190}{816233},$$

we obtain

$$\alpha_3 = 2(s_2 + s_6 + s_9 + s_{10} + s_{16} + s_{17} + s_{18}) + 4s_{11} = \frac{677498}{816233} \approx 0.830030151.$$

By Theorem 2, taking $\varepsilon = \alpha_3 - 0.83 > 0$, for n large enough we have almost surely that $M(n) \geq M_3(n) \geq 0.83n$. It can be noted in Table 1 that in practice this lower bound is satisfied.

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e_i	elem. of e_i	$f(e_i)$	neighbors of e_i
e_1		0	$(e_2, 1/2), (e_3, 1/2)$
e_2		2	$(e_1, 1)$
e_3		0	$(e_4, 1/6), (e_5, 1/6), (e_6, 1/3), (e_7, 1/6), (e_8, 1/6)$
e_4		0	$(e_4, 1/8), (e_5, 1/8), (e_6, 3/8), (e_7, 1/8), (e_9, 1/4)$
e_5		0	$(e_{10}, 3/4), (e_{11}, 1/4)$
e_6		2	$(e_2, 1/4), (e_{12}, 1/8), (e_{13}, 1/8), (e_{14}, 1/4), (e_{15}, 1/4)$
e_7		0	$(e_{10}, 3/4), (e_{11}, 1/4)$
e_8		0	$(e_{10}, 3/4), (e_{16}, 1/4)$
e_9		2	$(e_2, 3/10), (e_{12}, 1/10), (e_{14}, 1/5), (e_{15}, 1/5), (e_{17}, 1/5)$
e_{10}		2	$(e_2, 1/2), (e_3, 1/2)$
e_{11}		4	$(e_1, 1)$
e_{12}		0	$(e_2, 1/2), (e_3, 3/10), (e_6, 1/5)$
e_{13}		0	$(e_2, 1/2), (e_3, 3/10), (e_6, 1/5)$
e_{14}		0	$(e_2, 3/10), (e_3, 1/2), (e_6, 1/5)$
e_{15}		0	$(e_2, 1/2), (e_3, 3/10), (e_6, 1/5)$
e_{16}		2	$(e_2, 1/5), (e_{12}, 1/10), (e_{13}, 1/10), (e_{14}, 1/10), (e_{15}, 3/10), (e_{18}, 1/5)$
e_{17}		2	$(e_1, 5/6), (e_2, 1/6)$
e_{18}		2	$(e_1, 5/6), (e_2, 1/6)$

Table 2: The 18 states of the Markov chain for $k = 3$. In the 2nd column we show the first component of e_i , and in the 3rd column the second component $f(e_i)$. In the last column we show the neighbor states of e_i as a list of tuples of the form $(e_j, P_{i,j})$, where $P_{i,j} = \text{Prob}(X_{\ell+1} = e_j \mid X_\ell = e_i) > 0$ is the transition probability from e_i to e_j .

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