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7 **Trait-based numerical classification of mesic and wet grasslands in Poland**

8

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22 **Running title**

23 Trait-based classification of grasslands

24

25 **Abstract**

26 *Questions:* What vegetation types can be distinguished on the basis of plant functional traits
27 using numerical classification? How do they match syntaxonomical units?

28 *Location:* Poland

29 *Methods:* 6985 vegetation plots representing mesic and wet grasslands (*Molinio-*
30 *Arrhenatheretea*, *Polygono-Poetea*) were retrieved from the Polish Vegetation Database.
31 Plant functional trait data were assembled from the LEDA and Clo-Pla databases for most
32 species occurring in the data set. Community-weighted mean for five traits were calculated
33 for each plot: specific leaf area, canopy height, seed mass, bud bank index and clonality
34 index. Plots were classified using Ward's method and iterative relocation based on silhouette
35 widths. The clusters were interpreted and characterized on the basis of species and trait
36 composition, functional diversity, functional redundancy, Ellenberg indicator values, and
37 geographical distribution. The similarity between the trait-based classification and the
38 syntaxonomical assignment of plots is evaluated both statistically and by expert knowledge.

39 *Results:* Twelve clusters were distinguished. The classification mirrored the main gradients
40 structuring grasslands in Poland, although, some vegetation types with the strong dominance
41 of functionally unique species appeared more distinct than they are treated in syntaxonomy.
42 Clusters did not differ significantly in functional diversity and redundancy. The differences of
43 clusters in species and trait composition and environmental background are discussed.

44 *Conclusion:* The application of trait data and numerical methods is a promising approach for
45 obtaining vegetation classifications. Such classifications can be in closer relationship with the
46 most important ecosystem processes than floristic classifications because communities
47 comprising different species but similar functional trait distribution are not separated. Trait-
48 based classifications match phytosociological units to a variable degree. Functional
49 uniqueness and variation of abundance determines how individual species influence the
50 delimitation of vegetation types using our approach.

51

52 **Keywords**

53 Vegetation classification, plant traits, functional ecology, grasslands, *Molinio-*
54 *Arrhenatheretea*, community-weighted mean, functional diversity, functional redundancy,
55 numerical classification

56

57 **Introduction**

58 Due to its central role in ecosystem processes, vegetation characteristics are frequently used
59 as general descriptors of ecosystems or habitat types for the purposes of nature conservation,
60 land-use planning, and landscape mapping. Vegetation-plot databases are widely used for
61 establishing classifications, very often with the application of statistical methods (De Cáceres
62 et al., 2015). Such databases contain tens or hundreds of thousands of species by site records
63 collected during the long history of phytosociology, often with additional data on vegetation
64 physiognomy, geographical location and environmental background (Dengler et al., 2011).
65 These data sources make it possible to answer questions about vegetation variation on scales
66 as broad as countries or continents. As the essential type of data which is recorded in these
67 plots with the highest consistency is species occurrence (and often some form of abundance),
68 most typically these analyses use species as variables and sites as objects. In consequence, the
69 classifications reflect patterns in species composition, together with all the possible
70 mechanisms which influence community assembly, including selection, speciation, dispersal,
71 and drift (Vellend, 2010). However, some of these processes, e.g. random drift, may not be
72 interesting from the viewpoint of the potential user of the classification. Limited dispersal of
73 species has strong consequences on classification results. If the sample includes areas with
74 different site history, which is a common situation, species may not have had enough time to
75 colonise all habitat patches which would have been suitable for them. In this case, a
76 classification based on species composition will reflect not only environmental gradients but
77 differences in regional species pools. When the geographical extent of the study is very large,
78 and the effect of site history is strong, it can become impossible to reach a vegetation
79 classification reflecting environmental gradients, which would be valid over the entire study
80 area. This might be a primary reason for the high level of idiosyncrasy in national vegetation
81 classifications. Patterns of speciation are mainly relevant on biogeographic scales in time and
82 space; although, in specific studies differences in the phylogenetic structure of communities
83 may be important (Lososová et al. 2015). Nevertheless, most vegetation classification studies
84 seek answers for questions about what types of communities exist, and how their occurrence
85 is related to environmental gradients, and ecosystem functions – i.e., the process of selection.
86 However, it is increasingly recognised that patterns in species identities are not always tightly
87 related with ecosystem properties, instead, traits of species are more relevant from this
88 perspective (Díaz & Cabido, 2001; Díaz et al., 2004). Species respond to biotic and abiotic
89 factors by their traits (response traits), as well as they form their environment by them (effect

90 traits; Lavorel & Garnier, 2002). Species which have similar traits may substitute each other
91 without significantly altering ecosystem functioning – a phenomenon called functional
92 redundancy (Hooper et al., 2005). Integrating the trait-based approach should improve the
93 relevance of vegetation classification with respect to ecosystem functioning, and enhance the
94 generalizability of results. Hérault and Honnay (2007) already presented a classification
95 where instead of species, groups of species sharing similar traits called ‘emergent groups’
96 were applied as variables. Hence, it was possible to differentiate two types of riverine forests
97 in Luxembourg different in life-form spectra, dispersal modes, and conservation relevance.
98 However, this study restricted its scope on a specific vegetation type of a rather narrow area,
99 while the most typical challenge of recent vegetation classification works is providing
100 relevant and generalizable results over broad sample coverage in space and along ecological
101 gradients.

102 The functional approach of ecology has been in an intensive research period now for more
103 than a decade, partially shifting the focus away from the study of patterns of species-level
104 composition and diversity (Carlow, 1987; Tilman et al., 1997; McGill, Enquist, Weiher, &
105 Westoby, 2006). A major outcome of this field is the emergence of plant trait databases,
106 which provide trait measurements for thousands of taxa and hundreds of traits (Kattge et al.,
107 2011). Technically they are readily connectable with vegetation-plot databases, providing
108 avenues for the same types of analyses of trait composition which has only appeared at the
109 level of species yet. A fundamental question of trait-based ecology is the distinction between
110 processes which impedes the co-existence of functionally similar organisms (i.e. functional
111 divergence) and those which promotes it (i.e. functional convergence; Lhotsky et al., 2016).
112 Competitive exclusion is known to increase divergence according to the theory of limiting
113 similarity, while environmental (or niche) filtering (and, for traits increasing competitive
114 vigour, also competitive exclusion) supports functional convergence (Weiher & Keddy,
115 1995). Considering the interest of vegetation classification studies in the response of
116 vegetation to environment, trait convergence should be a key phenomenon in the construction
117 of functionally relevant classifications. Moreover, competitive exclusion is subordinated to
118 environmental filtering according to the filter model of community assembly (Keddy 1992),
119 which is a likely reason why trait convergence is more frequently detected than divergence.

120 In this paper we present a trial for integrating the functional approach into the context of
121 vegetation classifications using multivariate statistical methods. Our aim is to classify semi-
122 natural grasslands of Poland in a way that resulting groups are relatively similar in their trait

123 composition with no respect to species composition. We discuss the environmental
124 background, trait composition, functional diversity, and redundancy of the clusters
125 distinguished. We assess the similarity between the trait-based classification and the
126 syntaxonomical system. We expect the resulting classification to be generalizable over the
127 entire study area, while showing a strong relationship with ecosystem processes.

128

129 **Materials and Methods**

130 9725 phytosociological relevés representing temperate semi-natural grasslands were retrieved
131 from the Polish Vegetation Database (Kački & Śliwiński, 2012; GIVD identifier: EU-PL-
132 001). In the syntaxonomical system according to Kački, Czarniecka and Swacha (2013), these
133 grasslands are classified to *Molinio-Arrhenatheretea* class and comprise three orders, called
134 *Potentillo-Polygonetalia* (temporarily flooded and heavily grazed and trampled vegetation on
135 nutrient-rich soils), *Arrhenatheretalia elatioris* (lowland and montane mesic grasslands), and
136 *Molinetalia caeruleae* (wet grasslands and tall-forb vegetation). We included into the data set
137 also *Polygono arenastri-Poetea annuae* with one order *Polygono arenastri-Poetalia annuae*
138 (therophyte-rich dwarf-herb vegetation of trampled habitats) because it strongly resembles
139 heavily trampled pastures of the *Cynosurion* alliance (*Arrhenatheretalia*). *Polygono-Poetea*
140 *annuae* was treated as part of the *Molinio-Arrhenatheretea* in previous syntaxonomical
141 overviews in Poland. Plot size was restricted to 10 to 100 m². Moss and lichen species were
142 removed from the data set due to their uneven data availability across plots. Where species
143 covers were recorded on ordinal scales (e.g. Braun-Blanquet scales, which is ca. 90% of all
144 plots) cover categories were transformed to their respective mid-point percentages using the
145 JUICE software (Tichý 2002). Relevés with >5% cover of trees and shrubs were excluded.
146 The data set was subjected to geographical stratification and heterogeneity constrained
147 random resampling (Knollová, Chytrý, Tichý, & Hájek, 2005; Lengyel, Chytrý, & Tichý,
148 2011) using Bray-Curtis index calculated on square-root transformed abundance data. Strata
149 were 6'×10' in size. From each stratum the number of plots to select was determined as $5 +$
150 $(n-5) \times d$, where n is the total number of plots in the stratum and d is the mean pairwise
151 dissimilarity among plots within the stratum. This method down-weighted the contribution of
152 oversampled areas only if their beta-diversity was low, while intensively sampled but diverse
153 regions kept their high share (Wiser & De Cáceres 2013). No resampling was done in strata
154 containing five or less plots. The stratified resampling reduced the data set to 6985 plots.

155 A key decision in any trait-based study is the selection of traits to involve in the analysis.
156 According to Westoby (1998) and Westoby, Falster, Moles, Vesk, and Wright (2002), the
157 traits of the leaf economics spectrum (Wright et al., 2004), the height, and the seed represent
158 the major dimensions of plant variability along the most typical ecological gradients (the so-
159 called leaf-height-seed or LHS system). Leaf traits, especially specific leaf area (SLA;
160 Cornelissen et al., 2003) are in close connection with resource acquisition and relative growth
161 rate, and thus related to the productivity of the habitat (Wilson & Tilman, 1993; Lhotsky et
162 al., 2016). Canopy height is positively correlated with the ability to outcompete other species
163 in productive habitats, where light is the major limiting factor (e.g. Borer et al., 2014), while
164 seed mass corresponds to reproduction strategy (e.g. Moles & Westoby, 2004). Besides LHS
165 traits, there is growing evidence that clonal growth and bud bank are important in the
166 adaptation of plants to regular disturbance, hence they have central role in the response of
167 herbaceous vegetation to environment (Klimešová, Tackenberg, & Herben, 2016). Clonal
168 growth enables plants to avoid disturbance, while bud bank is key in regeneration after minor
169 damages. Five plant traits were included in the analysis, which can be regarded as the
170 response traits that play the most fundamental role in the adaptation of plants to the
171 environment and management regime. Specific leaf area (SLA), canopy height and seed mass
172 were retrieved from the LEDA database (Kleyer et al., 2008). The ‘bud bank’ and the
173 ‘clonality’ index were introduced according to Johansson, Cousins and Eriksson (2011) and
174 E.-Vojtkó et al. (2016). Bud bank index is the rank sum of above- and belowground bud bank
175 categories, while clonality index is the rank sum of lateral spread and total number of
176 offspring per parent categories. These data were obtained from the Clo-Pla database
177 (Klimešová, Danihelka, Chrtěk, de Bello, & Herben, 2017). All measurements were subjected
178 to a semi-automated outlier exclusion and averaging procedure for each species by traits.
179 First, the mean and standard deviation of all measurements from a given species and given
180 trait was calculated. Those measurements which differed by $>2*SD$ from the mean were
181 excluded. The remaining measurements were subjected to averaging weighted by the square-
182 root of the number of replications given for each measurement in the public database
183 (typically, the number of measured individuals for a given record). For species which lacked
184 no more than two out of the five trait values, Bayesian Hierarchical Matrix Factorization
185 (Schrodte et al., 2015) was used to fill the gaps in the trait table. Species with more than two
186 missing trait measurements were rejected from the analysis resulting in 885 species in the
187 final matrix. Plots where the relative cover of such rejected species was higher than 5% were
188 excluded before the stratified resampling. Species-level mean trait values were checked for

189 normality by quantile-quantile plots. Since LHS traits proved to be right-skewed, they were
190 log-transformed. Then, all traits were standardized to mean = 0 and standard deviation = 1.
191 These species-level standardized means were used for calculating community-weighted
192 means (CWM; Garnier et al., 2004), which is considered a satisfactory indicator of niche
193 filtering along large-scale environmental gradients (Kleyer et al., 2012; De Bello et al., 2013).
194 The CWM is constrained by the relative dominance of the most dominant trait value in the
195 plot; therefore we expect species forming monodominant vegetation stands to have a high
196 influence on the classification (De Bello, Lepš, Lavorel, & Moretti, 2007). This conforms the
197 mass ratio hypothesis by Grime (1998) stating that ecosystem functioning is mainly
198 determined by traits of the dominant species.

199 Hierarchical classification was carried out by Ward's agglomerative method (Podani, 2000)
200 on the matrix of CWM values. Ward's method relies on Euclidean distances in the trait space
201 between plots. The upper 20 hierarchical levels of the classification were evaluated by several
202 cluster validity indices (see Appendix S1); however, they suggested different numbers of
203 clusters as optimal making it impossible to decide on a single 'best' solution. To overcome
204 this, on the one hand, we considered also the biological interpretation of the clusters and the
205 resolution desired typically in such large-scale classifications. Hence, the dendrogram was cut
206 at a particular level and then it was improved using iterative relocation methods (Roberts
207 2015). Most iterative relocation methods proposed by Roberts (2015) are computationally
208 very demanding; therefore, we applied the REMOS2 algorithm (Lengyel, Roberts, & Botta-
209 Dukát 2019). This procedure uses the silhouette width index (Rousseeuw 1987) to identify
210 misclassified objects, which are then re-assigned to their closest neighbour cluster. After re-
211 assignment, silhouette widths are updated, and misclassified plots are relocated again to their
212 closest neighbour cluster, until the classification cannot be further optimized. This has
213 changed the assignment of 31.58% of all plots. However, we did not change the hierarchical
214 relations of the basic clusters.

215 Delimited clusters were interpreted as biologically relevant units using expert-based
216 knowledge and we attempted to find correspondence to already known and well-defined
217 vegetation units in the traditional (floristic) syntaxonomical approach (Kaçki et al., 2013). For
218 this purpose, phytosociological relevés were assigned to syntaxa at the class, order, and
219 alliance levels using their formal definitions. Relevés were classified to respective syntaxa
220 based on explicit definitions of vegetation units in the way that the relevé matched by the
221 definition of alliance must also match the definitions of the superior syntaxonomical units, i.e.

222 order and class. This classification system was created using combination of sociological
223 species groups (Bruehlheide, 1997), and total cover of individual species or group of species
224 (Kočí, Chytrý, & Tichý, 2003; Dengler et al., 2006; Landucci, Tichý, Šumberová, & Chytrý,
225 2015), and is part of an ongoing project of vegetation classification in Poland. The outcome of
226 this classification is presented in the shortened synoptic table (Appendix S2). Assignments at
227 order and alliance levels were compared with the trait-based classification using the
228 symmetric version of Goodman and Kruskal's lambda index (Goodman & Kruskal, 1954). To
229 assess the strength of similarity, the observed lambda values were compared to a reference
230 distribution obtained by re-calculating the index after permuting the trait-based classification
231 9999 times. In total, five tests were performed: 1) class-level assignment vs. trait-based
232 classification with random permutation; 2) order-level assignment vs. trait-based
233 classification with random permutation; 3) order-level assignment vs. trait-based
234 classification with restricted random permutation using class-level assignment as strata; 4)
235 alliance-level assignment vs. trait-based classification with random permutation; 5) alliance-
236 level assignment vs. trait-based classification with restricted random permutation using order-
237 level assignment as strata. We report the P-values of the null hypothesis stating that the
238 similarity between the trait-based classification and the syntaxonomical assignment is as high
239 as we can observe due to chance. We also show the standardized effect sizes of the observed
240 values calculated by probit transformation (Botta-Dukát, 2018).

241 For each level of the hierarchical classification until reaching the level of basic clusters and
242 for each trait a Wilcoxon test was carried out to test the difference between the two clusters to
243 be merged in the respective fusion. Two-tailed P-values were calculated by using permutation
244 tests (Hothorn, Hornik, van de Wiel, & Zeileis, 2006). Bonferroni-corrected P-values and
245 standardized test statistics (W_{st}) were used for ranking traits by support to the tested fusions.

246 Distribution of CWMs across clusters are shown on 'boxes-and-whiskers' plots. Clusters are
247 also compared on the basis of functional vulnerability measured by Rao Q diversity index
248 (Botta-Dukát, 2005), and functional redundancy (Ricotta et al., 2016). We also provide
249 synoptic tables containing diagnostic, constant and dominant species, Ellenberg indicator
250 values (EIV; Ellenberg et al., 1992), and geographical distribution of the clusters as
251 Appendices S3-S5. Statistical comparison of clusters on the basis of variables dependent on
252 compositional information (e.g. diversity indices, aggregated species attributes) often result in
253 false positive tests due to the so-called 'similarity issue' (Zelený & Schaffers, 2012; Zelený,
254 2018) and other structural biases (Hawkins et al., 2017) if the same occurrence information

255 had been used for the definition of clusters. To avoid false conclusions we refrain from formal
256 statistical tests in the case of between-cluster comparisons and present only boxes-and-
257 whiskers plots.

258 To measure the influence of individual species on the trait-based (especially, CWM-based)
259 classification we used a simple formula, that we call ‘influence index’. That species affects
260 CWM of a single community (e.g., a single relevé) the most, which has high abundance and
261 highly different trait value from the other species in the community. If the variation of CWM
262 across several communities is examined, species with highly variable abundance and unique
263 trait values are supposed to be the most influential. Thus, we calculated the influence index
264 for the k th species as follows:

$$265 \quad I_k = \sqrt{D_{.k} \times SD(\mathbf{a}_k)}$$

266 where $D_{.k}$ is the distance of the k th species from the unweighted mean of trait values of all
267 species in the data set (‘uniqueness’), and $SD(\mathbf{a}_k)$ is the standard deviation of the abundance
268 vector of species k (‘variation of abundance’). Thus, the influence index is the geometric
269 mean of two components, functional uniqueness and variation of abundance. We recommend
270 to re-scale both $D_{.k}$ and $SD(\mathbf{a}_k)$ by division by the maximum, respectively.

271 Calculations were carried out by the R software (R Core Team, 2017) using *vegan* (Oksanen
272 et al., 2018), *cluster* (Maechler, Rousseeuw, Struyf, Hubert, & Horni, 2017), *FD* (Laliberté,
273 Legendre, & Shipley, 2014), *rapporotools* (Blagotić & Daróczy, 2015), and *coin* (Hothorn,
274 Hornik, van de Wiel, & Zeileis, 2008) packages. Nomenclature of plants follow the
275 Euro+Med PlantBase (last accessed on 27 Sep 2018), syntaxon names are according to Kaçki
276 et al. (2013).

277

278 **Results**

279 Most cluster validity measures indicated a peak value between 2 and 5, as well as another
280 peak near 10 clusters (Appendix S1). After considering non-formal criteria, we chose the level
281 of twelve basic clusters for the interpretation because it provided a reasonable compromise
282 between details and conciseness; however, coarser solutions can easily be assembled from this
283 fine-scale classification by merging low-level clusters according to the dendrogram fusions
284 (Fig. 1). Distribution of CWM values across clusters is shown on Fig. 2. Synoptic tables and

285 textual descriptions of the clusters are presented in Appendix S3, EIVs in Appendix S4, and
286 geographic distributions in Appendix S5.

287 Clusters 1 to 4 represent different kinds of grazed or frequently cut, often trampled grasslands.
288 Cluster 1 consisted of plots of heavily grazed and trampled, or frequently cut grasslands from
289 various soil types and moisture levels. This cluster is characterised by the lowest canopy
290 among all clusters, high SLA and high clonality index. Cluster 2 represents mostly mesic
291 grasslands on nutrient-poor and acidic soils which are mown or occasionally grazed. In this
292 cluster SLA is slightly above and canopy height is slightly below the sample-wise average,
293 while clonality is similarly high as in Cluster 1. Cluster 3 represents a small and distinct group
294 of relevés dominated by *Agrostis stolonifera* and *Alopecurus geniculatus*. They occur on
295 trampled, sometimes slightly alkaline, irregularly inundated habitats with nutrient-rich soils.
296 In this cluster SLA is highest, canopy is second lowest, seed mass and bud bank are the lowest
297 across all clusters, while clonality is highest among them. Similarly to Cluster 2, Cluster 4
298 contains plots mostly from extensively grazed mesic grasslands. SLA, seed mass and bud
299 bank values of this cluster are slightly above the average of all clusters, canopy height is
300 below average, clonality is intermediate.

301 Clusters 5 and 6 are two large, heterogeneous clusters containing several mesic and wet
302 meadow types from lowland to montane sites. They are characterized by above-average seed
303 mass and bud bank. Canopy height is lower in Cluster 5 than in Cluster 6.

304 Clusters 7 and 8 are two, distinct types with monodominant graminoid species with high
305 canopy. Cluster 7 contains common lowland and montane wet grasslands, mostly with the
306 dominance of *Scirpus sylvaticus*. This cluster has the highest canopy on average, highest bud
307 bank on average, low seed mass and high clonality. Cluster 8 contains mesic meadows of
308 ruderal character with the dominance of *Arrhenatherum elatius* mostly on post-arable lands
309 converted to grasslands. In this cluster SLA, canopy, seed mass and clonality are high.

310 From Cluster 9 to 12 herbaceous communities of mostly wet habitats are found. Cluster 9
311 contains a variety of wet and mesic communities including intermittently wet meadows with
312 *Molinia caerulea*, wet tall-forb vegetation, and nutrient-rich mesic meadows. This cluster has
313 high seed mass, high bud bank, low clonality index and intermediate values for the other two
314 traits. Cluster 10 contains montane meadows, and degraded wet meadows with *Deschampsia*
315 *caespitosa* and to lesser extent *Juncus* species. This cluster was characterised by low SLA and
316 clonality, high bud bank, and intermediate values for the other two traits. Cluster 11 contained

317 a variety of wet meadows with constant presence of tall forbs. This cluster had low SLA and
318 clonality, and high values for the other three traits. Cluster 12 comprised relevés dominated
319 by *Juncus* species, most frequently *Juncus effusus*, occasionally *J. subnodulosus* or *J.*
320 *conglomeratus*. These stands occur mostly on nutrient-poor, waterlogged and acidic soils,
321 which are sometimes managed by grazing. This cluster has high canopy and low values for all
322 the other traits.

323 The cross-tabulation of the trait-based and the syntaxonomical classification is shown on Tab.
324 1. Permutation tests with Goodman and Kruskal's lambda index rejected the null hypothesis
325 stating that trait-based classification and syntaxonomical assignment are as similar as
326 expected by chance alone. Observed lambda values were higher than any element of the
327 reference distribution using either the class-level (lambda = 0.007), order-level assignments
328 (lambda = 0.313) or the alliance-level assignments (lambda = 0.236). In all cases $P < 0.001$
329 which gave SES = 3.719 after probit-transformation. However, the matching between
330 syntaxonomical and trait-based classifications was not perfect. *Potentillion anserinae*
331 (Clusters 2 and 3), *Juncion effusi* (Clusters 10 and 12), and *Polygono-Poetalia* (Clusters 1 and
332 4) were the few syntaxa which concentrated on a relatively limited number of trait-based
333 clusters, while the majority of other units were more broadly dispersed across several clusters.

334 At the highest hierarchical level (i.e., two clusters), clonality showed a difference between the
335 merged clusters which was the most extreme not only among all traits at that level but also
336 across all levels ($W_{st} = 66.56$; Table 2). From the three-cluster level onwards, we could found
337 no difference of this magnitude; although, with <7 clusters all tests showed significant
338 difference between the merged clusters. The bud bank showed the second largest difference
339 on absolute scale at the four-cluster level ($W_{st} = -38.04$). Apart from those mentioned above,
340 we could recognize no pattern in the contribution of individual traits to the merging of
341 clusters.

342 With some minor inequalities attributable to the unbalanced distribution of relevés, all clusters
343 were distributed over almost the entire country, none of them was obviously restricted
344 geographically.

345 In terms of functional diversity, clusters showed high overlap (Fig. 3). The highest median
346 Rao Q was detected in Cluster 12, while the lowest in Cluster 1. Cluster 12 showed also the
347 lowest functional redundancy together with Cluster 3. The other clusters resembled each other
348 very much also in this aspect.

349 On Fig. 4 we show the distribution of species in the space of D_k and $SD(a_k)$ and Table 3
350 shows the ten species with the highest scores. Three of the first four species are those, which
351 form monodominant and distinct vegetation types (*Arrhenatherum elatius*, *Scirpus sylvaticus*,
352 *Juncus effusus*), while the rest species also occur as typical dominants of certain clusters
353 (Appendix S3).

354

355 **Discussion**

356 In our paper we present the numerical classification of plots representing semi-natural
357 grasslands of Poland, based on plant trait data, more specifically, on community-weighted
358 trait means of phytosociological relevés.

359 Using the emergent group approach, Hérault and Honnay (2007) showed that the involvement
360 of trait data into classification could provide typologies which reflect certain ecosystem
361 properties better than what would be achieved using only species composition. The main
362 difference between Hérault and Honnay's approach and ours lays in how we took into account
363 trait information. Hérault and Honnay classified species on the basis of their trait values into
364 'emergent groups', which were used as variables instead of species. The power of the
365 emergent group approach is that it accounts for functionally redundant species explicitly,
366 since emergent groups consist of species possessing the same trait syndrome and thus having
367 very similar ecological functions. On the other hand, classification of species into discrete
368 groups requires subjective decisions from the researcher regarding the clustering algorithm,
369 similarity measure, and number of emergent groups. Moreover, even objective algorithms
370 produce non-intuitive classifications due to methodological constraints, e.g. certain methods
371 tend to prefer clusters with specific size or shape. Our approach avoided this pitfall by using
372 trait information as continuous variables to calculate CWMs which were input for
373 classification.

374 We divided the sample into 12 clusters based on biological interpretability; although, several
375 cluster validity indices had higher values at lower numbers of clusters. It might suggest that
376 the trait-based classification approach recognized coarser vegetation units than we found
377 relevant and well separable. Nevertheless, classification studies are often aimed at providing
378 vegetation typologies at different hierarchical levels, enabling practitioners to choose the most
379 suitable resolution for a given application. The finer cluster resolution discussed here allows a
380 more detailed overview of the whole sample with reduced within-cluster heterogeneity;

381 however, for specific purposes it is still possible to merge lower-level clusters, e.g., according
382 to the fusions of the dendrogram. Therefore, our view of vegetation classification and
383 typology suggested here is flexible to a degree.

384 The trait-based classification mirrored the most significant gradients shaping grassland
385 vegetation of Poland, i.e. soil nutrient supply, soil moisture, and management. At the highest
386 classification level, mostly nutrient-rich and mesic types (Clusters 1 to 8, except Clusters 3
387 and 7) were separated from communities of nutrient-poor and wet habitats (Clusters 9 to 12).
388 Clusters 1 to 4 form a separate group at the four-cluster level. Their separation at high
389 hierarchical level is notable, since these trampled and grazed, highly specialized grasslands
390 include plots which differ in species composition very much but they are rather similar in
391 terms of physiognomy and traits with characteristically low canopy and high SLA.

392 Goodman and Kruskal's lambda with a permutation test rejected the independency between
393 the trait-based and the syntaxonomical classification. This is not very surprising given that
394 both formal definitions and CWM values rely on the species composition of relevés; however,
395 we were not able to design a formal test of similarity with higher practical relevance since
396 there is no standard threshold for 'tolerable difference' determining whether two
397 classifications can be considered the same or not. With expert-based evaluation of the clusters
398 we could point out several mismatches between the trait-based classification and the
399 syntaxonomical system. A striking example can be seen in form of clusters which were
400 dominated by functionally unique species, e.g. *Scirpus sylvaticus* (Cluster 7), *Arrhenatherum*
401 *elatius* (Cluster 8), or *Juncus* spp. (Cluster 12). These vegetation types are either
402 differentiated at the association (e.g., *Scirpetum sylvatici*) or alliance level (e.g., *Juncion*
403 *effusi*), or not differentiated unequivocally (e.g., grasslands dominated by *Arrhenatherum*
404 *elatius*) in the syntaxonomical system, while in the trait-based classification they appeared as
405 very distinct clusters standing alone sometimes even at high hierarchical levels. Obviously,
406 functionally unique and monodominant types also defined as separate syntaxa increase
407 matching between syntaxonomic and trait-based classification, while syntaxonomically
408 undefined types decrease it. Since monodominant communities are often species-poor, their
409 distinct occurrence in the trait-based classification might be viewed as an artefact attributable
410 to differences in species richness, considering that the more species are selected from the total
411 species pool, the less likely it is to obtain an extreme community-weighted mean.

412 Nevertheless, we consider differences in dominance structure as a relevant aspect of the
413 biological phenomenon we study which mirrors environmental stress, disturbance, or specific

414 site-history that should not be removed from the analysis. The influence index accurately
415 identified those species which appeared as dominants of certain clusters; therefore, we
416 recommend its application for estimating the influence of individual species on CWM-based
417 classifications. On the other hand, several alliances with more balanced dominance structure,
418 higher species richness, and higher functional similarity between species did not separate well
419 in the trait-based classification. For example, most meadow alliances, including
420 *Arrhenatherion* and *Polygono-Trisetion* in the *Arrhenatheretalia* order, and *Calthion*,
421 *Cnidion*, and *Molinion* in *Molinietalia*, similarly occurred in Clusters 5, 6, and 9. Considering
422 the five traits we selected for our analysis, there is a high functional overlap between these
423 meadow types despite being assigned to different orders in the syntaxonomical system.
424 Importantly, the inclusion of other traits may explain specific functional differences between
425 these alliances and orders.

426 Clonality index was the trait having the highest influence on the classification, which is in line
427 with the findings of Klimesová et al. (2008, 2016) and E.-Vojtkó et al. (2016). Bud bank also
428 seemed to have a relatively strong impact. There are several possible reasons for the
429 efficiency of vegetative traits in revealing patterns in herbaceous vegetation. One reason is
430 that grasslands in the temperate zone are usually maintained by some form of biomass
431 removal, typically grazing or mowing. Plants adapt to such disturbances through avoidance or
432 regeneration using clonal and bud bank traits (Klimešová et al., 2016); although, affecting
433 other traits due to developmental trade-offs (Rusch, Wilmann, Klimešová, & Evju, 2011;
434 Herben, Šerá, & Klimešová, 2015). Differences in the form and timing of management may
435 be at least as significant as abiotic variation among the vegetation types included in this
436 analysis (i.e., mesic and wet, semi-natural grasslands without extreme conditions in abiotic
437 environment). Another potential explanation for the high influence of vegetative traits is that
438 the ability of clonal growth as expressed on the relatively coarse scale applied in the Clo-Pla
439 database shows lower levels of intraspecific variation due to stronger phylogenetic constraints
440 and less measurements error. We consider all these explanations similarly likely, and agree
441 that clonal and bud bank traits should be given high attention in the study of functional
442 responses of vegetation to environmental and management gradients.

443 We did not find striking difference between clusters in terms of functional diversity and
444 redundancy. Only Cluster 12 showed higher functional diversity and lower redundancy than
445 the others in median values, which can be explained also by the functional uniqueness and
446 high dominance of *Juncus* species. However, this may not be a reliable indication of the

447 vulnerability or conservation importance of this vegetation type, since *Juncus*-dominated
448 stands, especially with *Juncus effusus*, are very common on nutrient-poor, disturbed or
449 successional wetlands.

450 We had to apply simplifications during our analyses which could have limited us in revealing
451 certain patterns. We retrieved trait data from LEDA and Clo-Pla database with no respect to
452 the geographical and environmental origin of the records, thus neglecting an amount of
453 variation in trait values attributable to population-level adaptation to local conditions. With
454 the application of community-weighted mean in the description of plot-level trait values, and
455 using their Euclidean distances as dissimilarity measure, we neglected the role of intraspecific
456 variation, despite growing evidence on its significant role in community assembly and
457 response to environmental gradients (Bolnick et al., 2011; Violle et al., 2012; Siefert et al.,
458 2015). Phylogenetic constraints may also bias the relationship between CWM values and an
459 environmental gradient (Duarte, Debastiani, Carlucci, & Diniz-Filho, 2018). Since there is no
460 obvious implementation of phylogenetic correction into a classification framework yet, we
461 neglected this effect. Despite the greatest and honourable efforts of database curators,
462 differences in measurement protocols or technicalities may have caused an amount of
463 variation between data sets coming from different providers. Nevertheless, we believe that
464 these sources of bias do not compromise our results at the scale of the classification discussed
465 here.

466 In any trait-based study, the choice of the trait determines all the potential results and
467 conclusions. We included LHS, clonal and bud bank traits with equal weight because there is
468 growing evidence of their ability to describe major dimensions of plant variability and
469 response of plants to environmental and management gradients (Westoby et al., 2002;
470 Klimešová et al., 2016). We believe that these five response traits describe the most important
471 vegetation gradients appropriately in the analysed data set. However, different sets of traits, or
472 different weights attributed to them, may have resulted in fundamentally different
473 classifications. For specific studies, it is straightforward to select traits which are relevant for
474 the ecosystem property under study.

475

476 **Conclusions**

477 We prepared a classification system of a broad vegetation unit, semi-natural mesic and wet
478 grasslands of Poland, relying on plot-based numerical classification of community-weighted

479 means of LHS and vegetative plant traits. The classification mirrors differences in
480 management, moisture, as well as types dominated by functionally unique species (*Scirpus*
481 *sylvaticus*, *Arrhenatherum elatius*, *Juncus* spp.). Among all traits, clonal index had the
482 strongest influence on the classification. Although, the matching between the trait-based and
483 the syntaxonomical classification was closer than the randomized references applied here, it
484 varied across vegetation types. Syntaxa with high dominance of functionally unique species,
485 typically occurring under more stressed environmental conditions or specific site-history,
486 appeared distinct also in the trait-based classification. In contrast, syntaxa with typically more
487 balanced dominance structure and higher functional overlap between species did not separate
488 well in the trait-based classification. Despite some discrepancies with the traditional species-
489 based classification approach, functional trait-based classification provides biologically
490 interpretable clusters. It must be, however, noted that our classification was performed on a
491 type of vegetation highly dependent on management type and intensity. Classification of less
492 disturbed vegetation types may bring sharper delimitation of vegetation units and different
493 importance of individual traits.

494

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499

500 **Author contributions**

501 A.L. set the idea, carried out data analysis and led the writing. G.S. and Z.K. prepared the data
502 and added interpretation to the results. Z.B.D. gave suggestions on the methodology. All
503 Authors critically revised the manuscript.

504

505 **Data accessibility**

506 Vegetation plot data are accessible from the Polish Vegetation Database (Kącki & Śliwiński,
507 2012; GIVD identifier: EU-PL-001), trait data are available from the LEDA (Kleyer et al.,
508 2008) and Clo-Pla (Klimešová et al., 2017) databases.

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709 **Supporting information**

710 **Appendix S1:** Evaluation of the hierarchical classification using six validity indices.

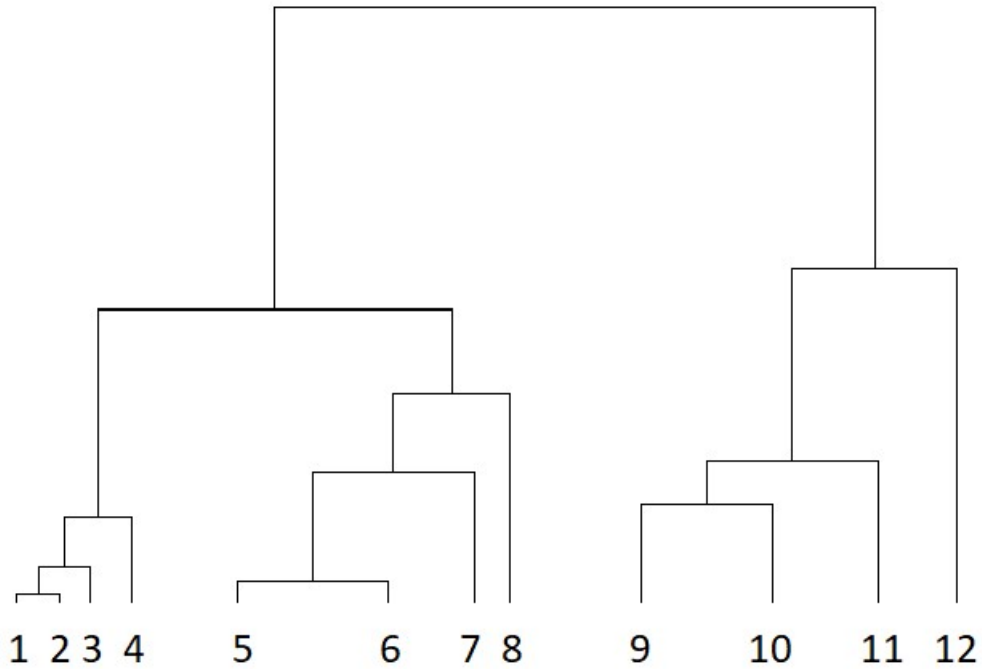
711 **Appendix S2:** Shortened synoptic tables of the syntaxonomic orders and alliances.

712 **Appendix S3:** Shortened synoptic table and textual description of the discussed clusters.

713 **Appendix S4:** Ellenberg indicator values of the discussed clusters.

714 **Appendix S5:** Geographical distribution of the discussed clusters.

715 **Figure 1.** Dendrogram of the upper 12 clusters in the hierarchy

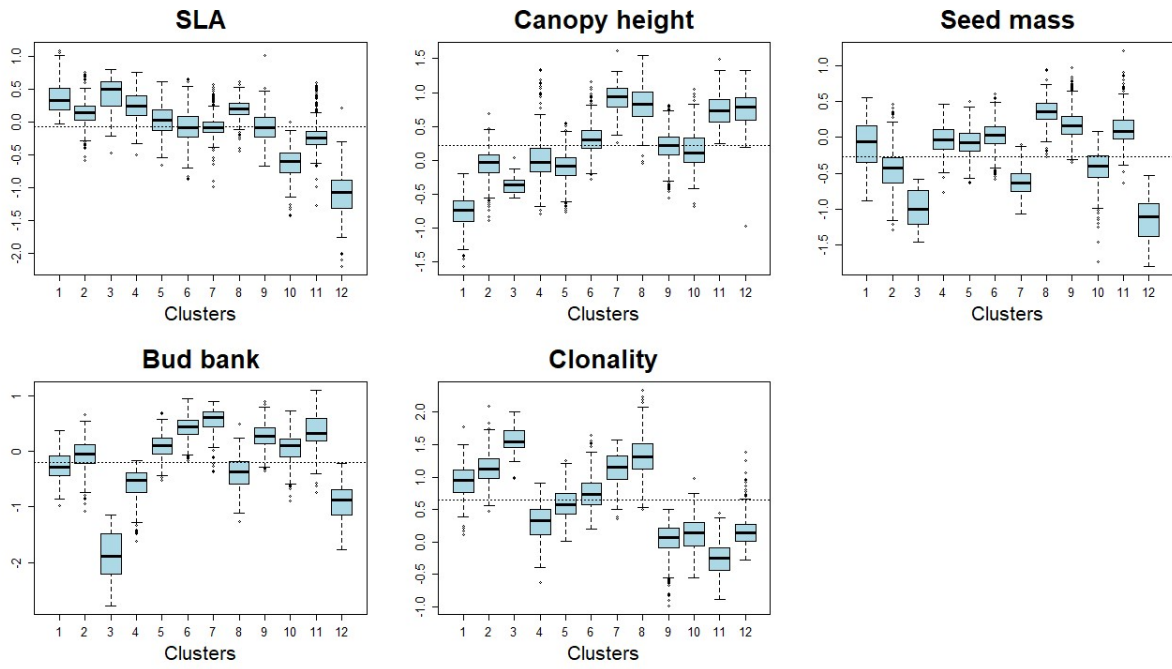


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719 **Figure 2.** Boxes-and-whiskers plots comparing community-weighted means of traits across
720 the twelve clusters. Boxes show upper and lower quartiles; whiskers show minimum and
721 maximum values; circles show outliers

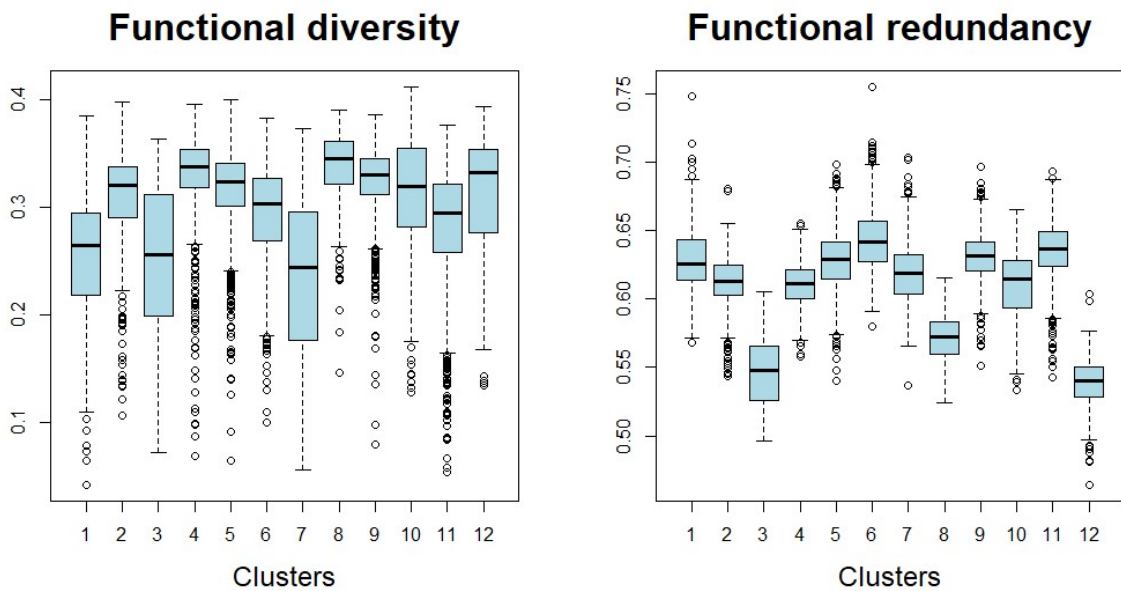


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724 **Figure 3.** Boxes-and-whiskers plots comparing Rao’s functional diversity and functional
 725 redundancy (Ricotta et al., 2016) across the twelve clusters. Boxes show upper and lower
 726 quartiles; whiskers show minimum and maximum values; circles show outliers

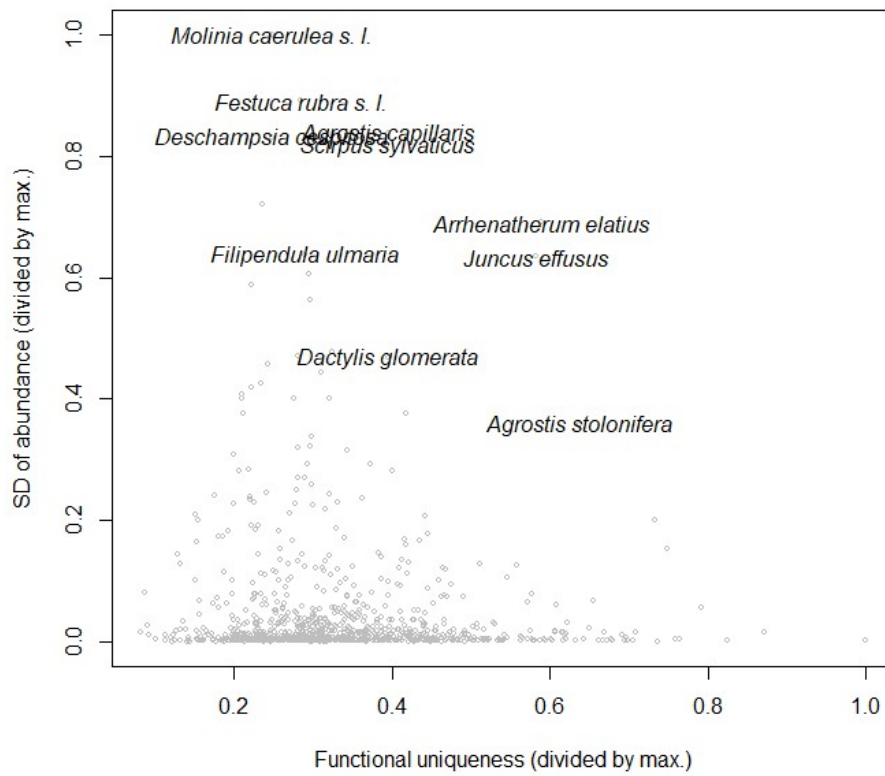
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730 **Figure 4.** Distribution of species in dimensions of functional uniqueness and SD of
731 abundance (both divided by maximum). Only names of the ten species with the highest
732 influence index values are shown



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734

735 **Table 1.** Cross-tabulation of the trait-based and the syntaxonomical classification

		Clusters of the trait-based classification													
CLASSES	ORDERS	ALLIANCES	1	2	3	4	5	6	7	8	9	10	11	12	
Syntaxonomical classification	Molinio-Arrhenatheretea	Arrhenatheretalia	<i>Arrhenatherion</i>	1	70	0	1	120	200	4	251	99	0	8	0
			<i>Cynosurion</i>	83	30	0	127	223	25	0	0	99	2	0	0
			<i>Poion alpinae</i>	0	0	0	1	0	0	0	0	1	0	0	0
			<i>Polygono-Trisetion</i>	0	11	0	0	34	24	0	0	24	1	2	0
			Unidentified at alliance level	9	287	0	39	446	452	1	11	190	7	4	0
		Molinetalia	<i>Calthion</i>	0	7	0	41	165	193	203	4	545	92	214	12
			<i>Cnidion</i>	0	0	0	7	35	55	1	0	49	28	4	14
			<i>Juncion effusi</i>	0	2	0	2	0	0	0	0	0	36	0	111
			<i>Molinion</i>	0	0	0	2	23	69	0	0	240	38	320	0
			Unidentified at alliance level	0	9	0	4	112	109	20	12	201	196	113	21
	Potentillo-Polygonetalia	<i>Potentillion anserinae</i>	1	11	11	0	1	0	0	0	0	2	0	0	
		Unidentified at alliance level	16	13	23	1	7	13	0	1	1	1	0	0	
	Unidentified at order level			14	41	0	13	153	260	11	10	33	7	2	2
	Polygono-Poetea	Polygono-Poetalia	Unidentified at alliance level	79	7	0	36	10	0	0	0	0	0	0	1
Dubious assignment			5	0	0	0	1	0	1	0	0	1	0	0	

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738 **Table 2.** Standardized test statistic of Wilcoxon tests of CWM-s between clusters to be
 739 merged in each fusion level of the hierarchical classification. Significance levels after
 740 Bonferroni adjustment (two-tailed tests): *** p<0.001, ** p<0.01, * p<0.05, ns = not
 741 significant

No. Clusters	SLA	Canopy height	Seed mass	Bud bank	Clonality
2	34.65***	-19.67***	-17.82***	-9.48***	66.56***
3	20.62***	-15.13***	20.93***	20.81***	-4.74***
4	20.53***	-29.81***	-19.41***	-38.04***	-6.83***
5	-15.69***	-22.37***	-22.22***	25.05***	-20.93***
6	3.37***	-36.72***	-12.05***	-17.44***	25.98***
7	2.22*	-25.42***	23.24***	-15.14***	-17.91***
8	28.88***	2.58**	31.22***	10.05***	-0.74 ^{ns}
9	3.19**	-15.86***	-15.15***	10.13***	24.96***

10	-5.5***	1.43 ^{ns}	8.52***	9.56***	-8.44***
11	18.81***	-17.4***	-0.79 ^{ns}	-30.07***	15.83***
12	8.47***	-18.19***	17.84***	-0.63 ^{ns}	-2.76**

742

743 **Table 3.** Species with the highest influence index

Species	Influence index
<i>Arrhenatherum elatius</i>	0.407
<i>Juncus effusus</i>	0.369
<i>Agrostis capillaris</i>	0.331
<i>Scirpus sylvaticus</i>	0.322
<i>Festuca rubra</i> s. l.	0.253
<i>Molinia caerulea</i> s. l.	0.248
<i>Agrostis stolonifera</i>	0.229
<i>Deschampsia cespitosa</i>	0.207
<i>Dactylis glomerata</i>	0.186
<i>Filipendula ulmaria</i>	0.184

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