The extent and specificity of relative age effects on mental health and functioning in

early adolescence

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Abstract

Background. Although extensive evidence indicates that being younger within a school cohort is associated with poorer academic functioning, much less is known about such relative-age effects (RAEs) for mental health.

Methods. Data from 23,378 11-13 year olds attending state-maintained secondary schools in England were analysed to investigate RAEs on mental health measured using the Strengths and Difficulties Questionnaire. Participants were grouped into oldest, middle and youngest thirds of their academic year based on their month of birth relative to their cohort. Hierarchical linear regression analysis evaluated RAEs and gender- or deprivation-related moderation of such effects.

Results. Relatively younger adolescents had significantly more emotional symptoms and peer problems compared to relatively older individuals in a year group, though effect sizes were small. These effects were not moderated by gender or deprivation. Impact of mental health difficulties on other aspects of functioning were also greater among relatively younger children. Larger RAEs are observed in the younger cohort (11-12 years) compared to in the 12-13 year olds, thereby indicating that RAEs might attenuate with age.

Conclusion. Being relatively younger than classmates is associated with increased internalising symptoms, poorer peer relationships, and higher impact of mental health difficulties on functioning at school and home. The findings support wider inclusion of relative age in understanding mental health difficulties and its inclusion as a potential risk factor in studies investigating the development of psychopathology, especially for internalising symptoms. Possible mechanisms of the effects detected are discussed.

Key words. Age within cohort, relative age, well-being, children, school

Implications and Contributions

This study demonstrates RAEs on mental health for emotional difficulties and peer problems, which are not moderated by gender or deprivation, but which appear to undermine other aspects of functioning. Though RAEs are small, from the perspective of multiple risks they merit consideration as a contributing factor in the development of internalising symptoms.

Introduction

Some children in a school year group are older than their classmates and some younger. This is the result of the use of calendar month cut-offs for the selection of cohorts into a year group. For instance, 1st September is used as the cut-off for school entry in England, meaning that those born in late August will be nearly one year younger than many of their classmates born in September the previous year. The term relative age effects (RAEs) is one of several used to describe the bias that disadvantages the youngest children in a given cohort. Whilst the effects of relative age on learning and school performance have been extensively researched (1-3), with evidence indicating that younger children are more at risk of poorer grades and of being identified as having special educational needs such as learning disorders (4, 5), much less RAE research has addressed mental health outcomes.

Only five investigations have focused on RAEs in mental-health-related outcomes. An early study valuating the impact of relative age on youth suicide, found that younger Canadian students within the cohort were more likely to have committed suicide during adolescence (6). A second investigation examined referral to psychological services in Northern Ireland, with results showing that younger children within the cohort were over-represented in referrals to psychological services (7). Two studies in British children aged 5-15 years (8) and 3-13 years (9) researched the impact of relative age on the Strengths and

Difficulties Questionnaire (SDQ) total difficulties score observing that younger children within cohorts had more difficulties overall. Their analyses also indicate that the negative impact of relative age decreases with age (9). Finally, Lien et al (10) documented RAEs on specific domains of mental ill-health at the end of secondary schooling (age 15-16 years) in a Norwegian sample and they found younger children manifesting fewer emotional difficulties and greater peer problems. Consideration of moderating effects of gender revealed, however, that whereas younger girls had fewer depressive symptoms than their older classmates, younger boys had more peer problems (10). No effects were observed for externalising symptoms, hyperactivity and pro-social behaviour sub-scales, highlighting the relevance of investigating whether relative age is associated with specific types of psychopathology symptoms (10).

A focus on the moderating role of gender—and also deprivation – in the Lien et al. work builds on such a focus in prior RAE research on academic outcomes (11-13), some of which highlights RAEs in educational outcomes moderated by both gender (e.g. RAEs larger in boys (10)) and deprivation (e.g. RAEs larger in higher SES (10)). Even though the effect of gender and deprivation on mental health is well documented (14), it remains unclear whether these factors moderate RAEs in the case of mental health. Thus, the research reported herein investigates the moderating effects of gender and deprivation on mental-health-related RAEs in early adolescence.

Cognitive and mental health RAEs indicate that the negative effect of being relatively younger within cohort decreases as children get older (9, 15). Therefore, even within the narrow age range examined in this report, we assess RAEs separately in two consecutive cohorts to assess whether they attenuate with age.

Studying mental-health-related RAEs is important because mental health difficulties can impact on many areas of life. Lien et al., (10) highlighted the importance of considering

downstream or knock-on effects of mental health RAEs on aspects of childhood and adolescent functioning, such as home life, classroom behaviour and friendships. Such potential derivative effects of mental-health-related RAEs are thus also a focus of the research reported herein.

Thus, the current research evaluates RAEs on five domains of childhood mental health (hyperactivity, conduct, emotional and peer problems, as well as, prosocial behaviour) and on the self-reported impact of these mental health difficulties on functioning in a large school-based sample of 11-13 years old secondary school students in England. Additionally, gender and deprivation are examined as potential moderators and RAEs are investigated separately by cohort. Based on work already cited, it was predicted that younger adolescents in a cohort would experience greater total difficulties than other children, though no specific predictions were advanced with regard to specific problem areas given the limited existing literature on RAEs and specific types of difficulties.

Method

Participants

Students from Year 7 (ages 11-12 years) and Year 8 (ages 12-13 years) in the English school system completed questionnaires as part of a wider study of mental health in schools(16). A total of 23,477 (73% response rate) students from 210 secondary schools participated in the wider study. Non-participation was mainly due to absenteeism followed by non-consent (1.25%). A small proportion (0.4%, n=93) of students were excluded from the current report, as they were outliers in terms of year or month of birth and the year group of their school (born a year or two earlier or later than the rest of their cohort). A further five individuals were excluded because they did not complete sufficient items in the measures used in this research.

The remaining 23,379 participants (Year 7 *N*=15362, Year 8 *N*=8017), comprising the analysis sample were born between September 1996 and August 1998 (age range 11.25-13.17, mean age=12.05,SD=.56). Just about half, 50.4% (N=11780), were female. The majority, 76.3%, were classified as being White followed by 9% Asian, 5.9% Black, 3.9% mixed ethnic background, 1.1% other ethnic groups and for 3.9% ethnicity information was unavailable. 19.2% (N=4489) were eligible for free school meals - a proxy for deprivation (17). Relative to the country (16), the analysis sample had higher proportions of ethnic minorities (24% vs. 18%) and individuals eligible for free school meals (19% vs. 13%).

Procedure

Computer-based surveys were completed by pupils within the normal school day. All students in a particular Year-group (7 or 8) in the school were eligible to participate. Information was sent to parents prior to data collection giving parents the opportunity to opt their child out from participating in the study. Additionally, participants had the research explained to them (in writing and orally), and were afforded the opportunity to decline participation. Student demographic information was obtained from linking records to the National Pupil Database, which is a nationally held dataset with school-related data on all pupils in England. The institutional research ethics committee of University College London reviewed the study and approved the data collection.

Measures

Relative age. Information about the month and year of birth for all participants was available via school records through the national pupil database. Based on their month and year of birth, they were divided into three relative age groupings, following the strategy of evaluating RAEs used by others (8, 10): 1) Oldest: born September to December (N= 7837), 2) Middle: born January to April (N= 7676), and 3) Youngest: born May to August (N= 7866).

Mental health. The Strengths and Difficulties Questionnaire (SDQ) (18), is used here as a measure of mental health and related difficulties (e.g., peer problems). This measure was selected as it is a widely used in research on young people and affords comparisons with previous RAE studies that relied on this measure (8, 10). The SDQ is a 25-item self-report measure consisting of five 5-item sub-scales: hyperactivity, emotional symptoms, conduct problems, peer problems and prosocial behaviour. In the current sample alpha estimates were hyperactivity .73, emotional symptoms .72, conduct problems .66, peer problems .61 and prosocial behaviour .69; indicating that many of the scales have poor internal consistency, which is not an uncommon finding with this measure (19). The four difficulty scales (emotional, peer, conduct and hyperactivity) are summed to create a total difficulties score. The scale have threshold scores to indicate clinical levels of problems (scoring and threshold details are available on http://www.sdqinfo.org).

Impact. The impact supplement of the SDQ assesses whether mental health problems affect other areas of the young person's functioning. The supplement starts with the question 'Overall, do you think that you have difficulties in one or more of the following areas: emotions, concentration, behaviour or being able to get on with other people?' and in the case of an affirmative response, participants answer five items (one concerning distress and four concerning impact on home-life, friendships, classroom learning and leisure activities) which are summed to create a total impact score.. A negative response to the initial item about presence vs. absence of difficulties results in a total impact score of 0. In the current sample 187 participants (0.8% of total sample) did not complete the items of the impact supplement.

Gender and socio-economic status (SES). Gender was recorded and coded 0=male, 1=female. As in other school-based research (17), socio-economic deprivation was coded in terms of eligibility for free school meals (0=not eligible for FSM, 1= eligible).

Analytic approach

First, descriptive analysis present demographic data across the three relative age groups and preliminary analysis to examine group differences in socio-demographic characteristics is carried out. Mean scores and proportions above the clinical threshold are presented for each of the mental health sub-domains, total difficulties and impact for each relative age grouping. To examine the contribution of relative age as a predictor of mental health difficulties, *hierarchical multiple regression models* were constructed to examine these effects across all outcomes (the mental health sub-scales, total difficulties score and impact score). In each model gender and SES (and their interaction) were included in the first step of analysis, followed by relative age in the second step (dummy coded with the oldest third as the reference category). In the third and final step the interactions between gender and relative age and SES and relative age were included to investigate possible moderation effects. In addition, whenever significant RAEs were detected, secondary analyses were conducted separately by year-group (Year 7 or 8) to illuminate any effect of age in moderating RAE effects.

Results

Descriptive statistics for the study variables are included in Table 1. Evaluation of group differences on key socio-demographic characteristics of interest in this study revealed no significant differences in gender proportions (χ^2 (2)=1.43, p=.49) and deprivation (proportion eligible for FSM, χ^2 (2)=1.22, p=.54).

Results of the hierarchical linear regressions, presented in Table 2, revealed significant main effects of gender and deprivation in the case of all outcomes. Specifically, girls scored higher than boys on emotional symptoms and prosocial behaviour and lower than boys on hyperactivity, conduct and peer problems. Deprivation was associated with more difficulties in all areas of mental health and with impact. With regard the primary focus of

inquiry, RAEs, these proved significantly related to the outcomes only in the case of emotional symptoms, peer problems, total difficulties and total impact. Comparing the youngest third to the oldest third of students indicates that the youngest third report significantly more symptoms for the emotional (β = .04, 95% CI .03-.06) and peer problem (β = .06, 95% CI .04-.07) sub-scales and total difficulties (β = .03, 95% CI .02-.05) and total impact (β = .03, 95% CI .01-.04). The standardised coefficients for relative age comparing the middle third to the oldest children in a year group indicate significantly more symptoms for both the emotional (β =.02, 95%CI .01-.04) and peer related difficulties (β = .04, 95% CI .02-.05) in the middle third of the cohort. Inclusion of interactions between relative age and gender and deprivation in Step 3 did not significantly improve the model for any of the mental health and impact indicators (all ΔR^2 = .000).

Analyses conducted separately in the two year groups to illuminate any agemoderated RAE effects are presented in Table 3. Results indicate that RAEs are more prominent in the younger group, aged 11-12 years, than in the cohort aged 12-13 years. The difference is especially striking for emotional symptoms where RAEs are entirely absent in the older cohort.

Discussion

Mental health difficulties are highly prevalent in young people. Understanding risk and protective factors associated with developing symptoms is of practical importance (20). The current study thus aimed to determine if being younger within school cohorts adversely affected sub-domains of mental health in early adolescence, a previously unexplored research question. Toward this end, we took advantage of a large school based data set of children aged 11-13 years in 210 secondary schools in England.

The RAE results on SDQ total difficulties are in line with previous findings (8, 9), who reported that total difficulties are highest in the youngest third of students in a cohort.

However, when it came to considering particular types of problems in the current inquiry, the mental-health disadvantage evident in younger children only emerged in the case of emotional symptoms and peer problems and not for conduct problems, hyperactivity or prosocial behaviour. Apparently, then, relative age only played a role in children's affect, thereby resulting in higher risk of internalizing symptoms, which includes symptoms of anxiety and depression. This impact of relative age could have longer term consequences as symptoms in childhood and adolescence are a pre-cursor for adult depression and psychopathology (21).

It is important to note that the detected RAEs proved to be small effects in terms of effect size, though at the population level still confer significant risk. The relevance is made clearer when we observe the proportions of young people with clinical levels of problems based on relative age groupings – based on the clinical threshold scores of the scales, there is a 1.2% difference between youngest and oldest students (12.2 vs. 11%) for emotional problems and a 4% (19.8 vs. 15.8%) for peer problems. The discrepancy is even more marked when considering only the younger cohort in the study, where a 2.2%, 4.7% and 3.5% difference in clinical cases between the youngest third and oldest third is observed for, respectively, emotional symptoms, peer problems and impact. The results support suggestions that youngest members of the cohort might have more difficulty being accepted by their peers (10), which is consistent with evidence indicating that they are much more likely to be bullied than other children (9). Conceivably, difficulties in reading, communication and attainment that are experienced by younger students, especially in the earlier years in school, might have an adverse impact on their socialization within their peer group (22, 23).

Analysis examining the potential role of gender and deprivation on RAEs failed to provide any evidence of moderation of RAEs by these factors, though as expected, greater deprivation predicted more difficulties across all outcomes. The null moderational results

suggest that the effects of relative age are distinct from gendered expressions of psychopathology and that relative age is an independent risk factor in predicting more internalising symptoms and peer problems for younger children within cohorts. These nongender-related results are notably different from those from the Norwegian study of 15-16 year olds which only found a significant negative effect of relative age on peer problems in the case of males (10). Recall, too, Lien and colleagues' (10) finding that only females showed RAEs for emotional symptoms— and then in a direction opposite to that predicted and documented herein (i.e. oldest females had more emotional symptoms than youngest females). The differences in results across these two studies could be accounted for by either the different age groups explored (11-13 vs. 15-16 years) or be indicative of cross-country differences in practice, whereby relative age mental health risks are more pronounced in England than Norway. If the former were the case, as children move into adolescence and both neurological and biological correlates become more gendered (24), gender-moderated RAEs as observed by Lien et al. could emerge in middle adolescence.

In terms of cohort differences, within the two consecutive cohorts that the current study covers, we observe that relative age effects are stronger in the younger cohort aged 11-12 years. This is in line with findings on attainment that suggest the strength of RAEs decrease with age (25). Developmentally this is not a surprising finding, as at younger ages the developmental differences that a year confers are more marked when compared to during adolescence. This may help to explain the results of Lien et al. (2005) who discerned virtually no RAEs for mental health among 15-16 year olds. It can also be expected from our and Lien et al.'s findings that RAEs might be more marked for internalising symptoms in even younger children than those included in these two studies. Even if this proves to be the case, it is possible that such age-related RAEs could be a function, to some extent, of developmental differences in symptoms.

The results on impact provide fresh insight into the possible mechanisms involved in the relative age disadvantage in a diverse range of domains. So far, due to the bulk of research on RAEs in the school context having focused on educational attainment and learning outcomes (1), explanations regarding such effects have tended to focus on ability grouping and special needs (26). Although the ability-grouping explanation has received some empirical support recently for learning outcomes (26), it is not clear how it could account for RAEs in the case of mental health and well-being. The current findings that those younger in cohorts report higher impact of their difficulties in functioning at home, in the classroom and with peers suggest that their ability to cope with their mental health difficulties is lower than that of their older peers. This seems possibly a function of less well developed coping mechanisms resulting from having experienced higher stress through childhood – a possible focus of future inquiry.

The results from the current study suggest many different mechanisms might be leading to RAEs in mental health. One is that problems with peers due to poorer cognitive and social skills in early school years might lead to developing emotional symptoms that impact on relationships in the classroom and home (27). It is also conceivable that self-esteem plays a role in a self-perpetuating cycle, and research demonstrates that relative younger age adversely affects self-esteem (28). Consider in this regard the possibility that relatively younger children feel less skilled than relatively older ones and that this undermines their self-esteem, which then makes them feel even less capable, with this negative feedback continuing over time. Alternatively, or additionally, it may be that relatively older children feel more skilled than their younger classmates and this leads them to experience enhanced self-esteem, which becomes part of a similar, even if opposite, self-perpetuating cycle. Longitudinal investigations are necessary to uncover the timeline and sequence of processes involved in developing mental-health-related RAEs. There is the

possibility that negative consequences beget more negative consequences which leads to a self-perpetuating cycle resulting in lifelong disadvantage (29). Not inconsistent with this view is evidence that the relative age disadvantage remains evident in post-compulsory education and unemployment in adulthood (30).

Study strengths and limitations

The current study used a large community-based sample and a widely used measure of mental health. Although we found that more of the youngest children within cohorts scored above clinical threshold than older children, the current study was not positioned to determine how this translates into higher risk of diagnosis or need of specialist mental health treatment. However, there is evidence that the SDQ is a fairly reliable public health indicator of clinical need (31), and existing research indicates that disproportionately more of the youngest children in a school cohort are likely to be receiving psychology services and intervention (7) and special educational needs assessment (12, 32).

The current study is also limited by the measures used, the cross-sectional design and the narrow age range that were the focus of inquiry. Further research would benefit from exploring RAEs in mental health, especially internalising symptoms and disorders, at different developmental stages to ascertain the effects of being younger in a cohort through childhood and adolescence. Investigating a wider range of moderators such as family factors and parenting might help understand those children at higher risk of experiencing RAEs. Additionally, investigation of RAEs on symptom development in longitudinal data will help illuminate the significance of RAEs as a risk factor for developing mental disorders.

Another limitation of the current study is the use of school-year based cut-offs, which although pertain to all children within schools, are not specific to mental health development (whereas RAEs in school attainment are linked with school cut-off dates and sporting achievement to sport cut-off dates). It is possible that setting-specific cut-offs such as for

sport teams, might affect the mental health of children involved in these activities, which future research might investigate.

Conclusion

Suggested strategies or interventions to prevent or minimise RAEs have mainly been based on effects found for academic attainment and take one of two broad forms: 1) adaptations to the school admission process and system overhauls (e.g., staggered school starting dates, children starting school on a particular birthday irrespective of the school year and month-of-birth based in-school grouping); and, 2) interventions to increase support to reduce or prevent the disadvantage faced by younger individuals in any given cohort (e.g., use of age-standardized tests, increased RAE awareness among teachers and educational psychologists, monitoring referral rates to psychiatric units etc.) (33, 34). There is limited evidence to suggest any of these strategies are effective. Clearly, understanding the domains in which RAEs operate, extent of the impact, and mechanisms underlying RAEs is required to inform the development of interventions and strategies (34). In regards to mental health difficulties this study begins to provide evidence that only certain domains of mental illhealth are affected for younger individuals in cohorts. Further research is necessary in this area to help understand the developmental patterns of RAEs and the mechanisms involved for internalising and peer problems. Longitudinal data analysis and investigation of potential mediators will be a useful next step to unpack mental health related RAEs in childhood and adolescence and their later life impacts.

The results of the current study highlight the small yet relevant role that relative age within cohorts might play in the development of young people's internalising symptoms and peer relationships. This is pertinent given for mental health multiple risk factors can have greater effect cumulatively, hence, even if effects of being younger in cohorts are small, they should be taken into account when considering who is most at risk. The findings support its

inclusion as a relevant risk factor more widely in studies of development of psychopathology and school based mental health intervention research.

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Table 1: Descriptive statistics of key variables by relative age grouping

		Oldest (Sept- Dec born)	Middle (Dec- April born)	Youngest (May-Aug born)	Total sample
Gender	% Female	50.9%	50.1%	50.1%	50.4%
SES	% FSM eligible	20%	19.9%	19.3%	19.7%
Hyperactivity	Scale score	3.84 (2.39)	3.82 (2.36)	3.84 (2.85)	3.83 (2.37)
	Clinical %	23.5%	22.9%	23.0%	23.1%
Conduct problems	Scale score	2.10 (1.96)	2.11 (1.95)	2.11 (1.94)	2.10 (1.95)
	Clinical %	22.3%	22.2%	22.5%	22.3%
Emotional Symptoms	Scale score	2.60 (2.20)	2.69 (2.21)	2.78 (2.24)	2.69 (2.22)
	Clinical %	11.0%	11.3%	12.2%	11.5%
Peer problems	Scale score	1.82 (1.76)	1.95 (1.83)	2.03 (1.84)	1.93 (1.81)
	Clinical %	15.8%	18.4%	19.8%	18%
Prosocial behaviour	Scale score	7.42 (1.94)	7.48 (1.92)	7.46 (1.91)	7.45 (1.92)
	Clinical %	17.0%	16.1%	16.7%	16.6%
Total difficulties	Scale score	10.36 (5.96)	10.57 (5.95)	10.76 (6.03)	10.56 (5.98)
	Clinical %	19.0%	20.3%	21.7%	20.3%
Total impact	Scale score	.78 (1.79)	.82 (1.82)	.88 (1.87)	.83 (1.83)
	Clinical %	24.3%	25.7%	26.9%	25.7%

SES- socio-economic status

Table 2: Hierarchical multiple regression analyses predicting mental health outcomes

	Hyperactivity		Conduct problems		Emotional symptoms		Peer problems		Prosocial behaviour		Total difficulties		Total impact	
Variable	ΔR^2	β	ΔR^2	β	ΔR^2	β	ΔR^2	β	ΔR^2	β	ΔR^2	β	ΔR^2	β
Step 1	.022**		.068**		.022**		.017**		.078**		.026**		.014**	
Gender		13**		22**		.14**		09**		.27**		10**		07**
SES		.08**		.15**		.05**		.09**		07**		.13**		.09**
Gender*SES		01		03*		.00		.00		.02		01		02
Step 2	.000		.000		.001**		.003**		.000		.001**		.001*	
Gender		13**		22**		.14**		09**		.27**		10**		07**
SES		.08**		.15**		.05**		.09**		07**		.13**		.09**
Gender*SES		01		03*		.00		.00		.02		01		02
Relative age (middle)		01		.00		.02*		.04**		.02		.02		.01
Relative age (youngest)		00		.00		.04**		.06**		.01		.03**		.03**
Step 3	.000		.000		.000		.000		.000		.000		.000	
Gender		13**		22**		.15**		09**		.27**		10**		05**
SES		.09**		.15**		.06**		.08**		09**		.13**		.10**
Gender*SES		01		03*		.00		.00		.02		01		02
Relative age (middle)		00		.00		.03*		.05**		.01		.03		.03*
Relative age (youngest)		00		.00		.04**		.05**		.01		.03*		.04**
Relative age (middle)*gender		.00		.00		01		02		01		01		02
Relative age (middle)*SES		00		01		02		00		.02		01		01
Relative age (youngest)*gender		.01		.01		01		00		01		.01		02
Relative age (youngest)*SES		01		00		.01		.02		.02		.00		01
Total R^2	.022**		.068**		.024**		.020**		.078**		.027**		.015**	

Note. *p<.01, **p<.001; SES- socio-economic status

Table 3. Regression analysis predicting relative age effects for emotional symptoms, peer problems and impact separately in the two cohorts

		Year 7 sample		Year 8 sample				
	Emotional symptoms	Peer problems	Impact	Emotional symptoms	Peer problems	Impact		
	B(SE)	B(SE)	B(SE)	B(SE)	B(SE)	B(SE)		
Gender	0.56** (0.04)	-0.33** (0.03)	-0.31** (0.03)	0.75** (0.05)	-0.31** (0.04)	-0.15** (0.04)		
SES	0.33** (0.04)	0.44** (0.04)	0.46** (0.04)	0.19* (0.06)	0.34** (0.05)	0.32** (0.05)		
Relative age (middle)	0.16** (0.04)	0.17** (0.04)	0.08 (0.04)	0.00 (0.06)	0.10 (0.05)	0.00 (0.04)		
Relative age (youngest)	0.26** (0.04)	0.28** (0.04)	0.16** (0.04)	0.09 (0.06)	0.12* (0.05)	0.01 (0.04)		

^{*}p<.01, **p<.001; SES- socio-economic status