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Spatial Depth-Based Methods for Functional Data

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Data**

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To Ilaria, Nunzia e Rino

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Abstract

In this thesis we deal with functional data, and in particular with the notion of functional depth. A functional depth is a measure that allows to order and rank the curves in a functional sample from the most to the least central curve. In functional data analysis (FDA), unlike in univariate statistics where \mathbb{R} provides a natural order criterion for observations, the ways how several existing functional depths rank curves differ among them. Moreover, there is no agreement about the existence of a best available functional depth. For these reasons among others, there is still ongoing research in the functional depth topic and this thesis intends to enhance the progress in this field of FDA.

As first contribution, we enlarge the number of available functional depths by introducing the kernelized functional spatial depth (KFSD). In the course of the dissertation, we show that KFSD is the result of a modification of an existing functional depth known as functional spatial depth (FSD). FSD falls into the category of global functional depths, which means that the FSD value of a given curve relative to a functional sample depends equally on the rest of the curves in the sample. However, first in the multivariate framework, where also the notion of depth is used, and then in FDA, several authors suggested that a local approach to the depth problem may result useful. Therefore, some local depths for which the depth value of a given observation depends more on close than distant observations have been proposed in the literature. Unlike FSD, KFSD falls in the category of local depths, and it can be interpreted as a local version of FSD. As the name of KFSD suggests, we achieve the transition from global to local proposing a kernel-type modification of FSD.

KFSD, as well as any functional depth, may result useful for several purposes. For in-

stance, using KFSD it is possible to identify the most central curve in a functional sample, that is, the KFSD-based sample median. Also, using the $p\%$ most central curves, we can draw a $p\%$ -central region ($0 < p < 100$). Another application is the computation of robust means such as the α -trimmed mean, $0 < \alpha < 1$, which consists in the functional mean calculated after deleting the proportion α of least central curves. The use of functional depths in FDA has gone beyond the previous examples and nowadays functional depths are also used to solve other types of problems. In particular, in this thesis we consider supervised functional classification and functional outlier detection, and we study and propose methods based on KFSD.

Our approach to both classification and outlier detection has a main feature: we are interested in scenarios where the solution of the problem is not extremely graphically clear. In more detail, in classification we focus on cases in which the different groups of curves are hardly recognizable looking at a graph, and we overlook problems where the classes of curves are easily graphically detectable. Similarly, we do not deal with outliers that are excessively distant from the rest of the curves, but we consider low magnitude, shape and partial outliers, which are harder to detect. We deal with this type of problems because in these challenging scenarios it is possible to appreciate important differences among both depths and methods, while these differences tend to be much smaller in easier problems.

Regarding classification, methods based on functional depths are already available. In this thesis we consider three existing depth-based procedures. For the first time, several functional depths (KFSD and six more depths) are employed to implement these depth-based techniques. The main result is that KFSD stands out among its competitors. Indeed, KFSD, when used together with one of the depth based methods, i.e., the within maximum depth procedure, shows the most stable and best performances along a simulation study that considers six different curve generating processes and for the classification of two real datasets. Therefore, the results supports the introduction of KFSD as a new functional depth.

For what concerns outlier detection, we also consider some existing depth-based procedures and the above-mentioned battery of functional depths. In addition, we propose three new methods exclusively designed for KFSD. They are all based on a desirable feature for a

functional depth, that is, a functional depth should assign a low depth value to an outlier. During our research, we have observed that KFSD is endowed with this feature. Moreover, thanks to its local approach, KFSD in general succeeds in ranking correctly outliers that do not stand out evidently in a graph. However, a low KFSD value is not enough to detect outliers, and it is necessary to have at disposal a threshold value for KFSD to distinguish between normal curves and outliers. Indeed, the three methods that we present provide alternative ways to choose a threshold for KFSD. The simulation study that we carry out for outlier detection is similarly extensive as in classification. Besides our proposals, we consider three existing depth-based methods and seven depths, and two techniques that do not use functional depths. The results of this second simulation study are also encouraging: the proposed KFSD-based methods are the only procedures that have good correct outlier detection performances in all the six scenarios and for the two contamination probabilities that we consider.

To summarize, in this thesis we will present a new local functional depth, KFSD, which will turn out to be a useful tool in supervised classification, when it used in conjunction with some existing depth-based methods, and in outlier detection, by means of some new procedures that we will also present in this work.

Resumen

El tema de esta tesis es el análisis de datos funcionales, y en particular de la noción de profundidad funcional. Una medida de profundidad funcional permite ordenar las curvas de una muestra funcional de la más central a la menos central. Al contrario de lo que ocurre en \mathbb{R} donde existe una forma natural de ordenar las observaciones, en el análisis de datos funcionales (FDA) no existe una forma única de ordenar las curvas, y por tanto las diferentes profundidades funcionales existentes ordenan las curvas de distintas formas. Además, no existe un acuerdo sobre la existencia de una profundidad funcional mejor para todas las situaciones entre las disponibles. Por estas razones, entre otras, el tema de la noción de profundidad funcional es todavía un área de estudio de investigación activa, y esta tesis se propone colaborar en los avances en este campo de FDA.

Como primera contribución, en esta tesis se amplía el número de profundidades funcionales disponibles mediante la introducción de la profundidad espacial funcional kernelizada (KFSD). A lo largo de este trabajo, se muestra que KFSD es el resultado de una modificación de una profundidad funcional existente conocida como profundidad espacial funcional (FSD). FSD se puede englobar dentro de la categoría de las profundidades funcionales globales, lo que significa que el valor de FSD para una curva dada, en relación con una muestra funcional, depende igualmente del resto de las curvas en la muestra. Sin embargo, como en el contexto multivariante, donde también se utiliza el concepto de profundidad, varios autores han sugerido que un enfoque local para la definición de una profundidad puede resultar útil también en FDA. Por este motivo, en la literatura se han propuesto algunas profundidades locales para las que el valor de la profundidad de una observación depende más de las ob-

servaciones cercanas que de las distantes. A diferencia de FSD, KFSD se puede clasificar en la categoría de las profundidades locales, y puede ser interpretada como una versión local de FSD. Como el nombre de KFSD sugiere, la transición de lo global a lo local se logrará mediante una modificación de FSD basada en el uso de los kernels.

KFSD, así como cualquier otra profundidad funcional, puede resultar útil para varios propósitos en el ámbito del análisis estadístico de datos. Por ejemplo, usando KFSD es posible identificar la curva más central en una muestra funcional, es decir, la mediana de la muestra según KFSD. Además, utilizando el $p\%$ de las curvas centrales, es posible definir la $p\%$ -región central ($0 < p\% < 100$). Otra aplicación es el cálculo de medias robustas, como por ejemplo la α -media truncada, con $0 < \alpha < 1$, que consiste en la media funcional calculada sin considerar la proporción α de las curvas menos centrales. El uso de las profundidades funcionales en FDA ha ido más allá de los ejemplos anteriores, y en la actualidad las profundidades funcionales también se utilizan para resolver otros tipos de problemas. En particular, en esta tesis se consideran la clasificación supervisada funcional y la detección de curvas atípicas, y se estudian y proponen métodos basados en KFSD.

El enfoque que se presenta en esta tesis en clasificación y detección de atípicos tiene una característica principal: el foco del trabajo está puesto en escenarios en los que la solución del problema no resulta muy clara gráficamente. Específicamente, en el apartado de clasificación se consideran casos en los que los diferentes grupos de curvas son apenas reconocibles mirando un gráfico, mientras que no se consideran problemas donde las clases de las curvas son fácilmente detectables gráficamente. De manera similar, no está entre nuestros objetivos detectar curvas atípicas que están excesivamente alejadas gráficamente del resto de las curvas, y por el contrario se consideran atípicos de baja magnitud, de forma y atípicos parciales, que son más difíciles de detectar con los procedimientos que ya existen en la literatura. En este sentido, se pondrá en evidencia que en este tipo de problemas existen diferencias sustanciales entre las profundidades y los métodos de análisis, mientras que estas diferencias tienden a ser menores en problemas más sencillos o visualmente más evidentes.

En relación con el problema de clasificación funcional, existen en la literatura métodos basa-

dos en el uso de las profundidades funcionales. En esta tesis se consideran tres procedimientos de este tipo, y por primera vez se combinan con varias profundidades funcionales (KFSD y seis más) con el objetivo de establecer comparativas entre métodos y/o profundidades con los mismos escenarios. El resultado principal que se observa es que KFSD se destaca entre sus competidores. De hecho, KFSD, cuando se utiliza junto a uno de los métodos conocido como el procedimiento de profundidad máxima en los grupos, muestra los resultados mejores y más estables a lo largo de un estudio de simulación que considera seis procesos diferentes para generar las curvas, así como en la clasificación de dos conjuntos de datos reales. Por lo tanto, los resultados obtenidos sustentan la introducción de KFSD como nueva profundidad funcional.

Por lo que se refiere a la detección de curvas atípicas, también se consideran algunos procedimientos ya existentes basados en el uso de la noción de profundidad y el grupo de siete profundidades mencionado arriba. Además, se proponen tres nuevos métodos diseñados exclusivamente para KFSD. Todos ellos se basan en una característica deseable en una profundidad funcional, es decir, que ésta asigne un valor de profundidad baja a una curva atípica. Durante nuestra investigación, se ha observado que KFSD posee esta característica. Además, gracias a su enfoque local, KFSD es en general capaz de ordenar correctamente los atípicos que no se destacan claramente en un gráfico. Sin embargo, un valor bajo de KFSD no es suficiente para detectar curvas atípicas, y es necesario tener a disposición un valor umbral para KFSD para distinguir entre curvas normales y atípicas. De hecho, los tres métodos que se presentan ofrecen formas alternativas para elegir un umbral para KFSD. Desde un punto de vista metodológico, estos procedimientos están respaldados por resultados teóricos de corte probabilísticos. El estudio de simulación que se lleva a cabo para la detección de atípicos es igualmente extenso como en el caso de clasificación. Además de nuestras propuestas, se consideran tres métodos existentes que están basados en el uso de profundidades funcionales y dos técnicas que no utilizan profundidades funcionales. Los resultados de este segundo estudio de simulación son también positivos: los métodos basados en KFSD que se proponen en esta tesis resultan ser los procedimientos que detectan mejor los atípicos para un conjunto de

seis escenarios simulados y para las dos probabilidades de contaminación que se consideran.

En resumen, en esta tesis se presenta una nueva profundidad funcional local, KFSD, que resulta ser una herramienta útil en clasificación supervisada cuando se utiliza conjuntamente con algunos métodos basados en el uso de profundidades, y en la detección de curvas atípicas por medio de algunos nuevos procedimientos que también se presentan en este trabajo.

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*La poesia non è di chi la scrive,
è di chi gli serve!
(Mario Ruoppolo to Pablo Neruda in Il Postino)*

Chapter 1

Introduction

The technological advances of the last decades in fields such as chemometrics, engineering, finance, growth analysis or medicine among others have allowed to observe random samples of curves. In these cases it is common to assume that the curves have been generated by a stochastic function and to refer to them as functional data. More precisely, a functional datum y is an observation of a functional random variable $Y \in \mathbb{H}$, where \mathbb{H} is an infinite-dimensional (functional) space. Alternatively, Y can also be interpreted as a stochastic process $\{Y(s), s \in I\}$, where I is an interval in \mathbb{R} , and the functional datum expressed as $y(s)$.

In practice, a functional datum $y(s)$ is always observed at a finite number of evaluation points, that is, in the form of $y(\mathbf{s}) = (y(s_1), \dots, y(s_m))$, where $\mathbf{s} = (s_1, \dots, s_m)$ is the set of domain points where $y(s)$ has been measured. However, even if $y(\mathbf{s})$ is a vector, there are at least three reasons why it is harmful to treat functional data as multivariate data.

1. The sets of domain points where the functional data of a sample are measured may differ in size and/or elements among them. Moreover, for each observation the order of the vector \mathbf{s} contains an important amount of information and permutations of its elements are not allowed, unlike in the multivariate case.
2. Any curve is a realization of a stochastic function with a certain dependence structure. Consequently, functional data are often autocorrelated. Therefore, it is hard to analyze

functional data by means of standard multivariate procedures because they usually fail in presence of autocorrelation.

3. Functional samples may contain less curves than evaluation points or, in multivariate jargon, less rows than columns. It is well known that matrices with more columns than rows are hardly handled by multivariate techniques.

As a response to the difficulty described above of multivariate data analysis to deal with functional data, the area of statistics known as functional data analysis (FDA) has arisen in recent years. Two seminal and complementary books about FDA, one parametric and the other nonparametric, are [Ramsay and Silverman \(2005\)](#) and [Ferraty and Vieu \(2006\)](#), respectively. More recently, [Horváth and Kokoszka \(2012\)](#) focused on inference and asymptotic theory for functional data, whereas [Cuevas \(2014\)](#) presented a partial overview of the state of art in FDA theory. To give an idea of the kind of data FDA deals with, we show in [Figure 1.1](#) three real datasets that are considered in this thesis. They are:

- growth curves of 54 heights of girls measured at a common discretized set of 31 nonequidistant ages between 1 and 18 years ([Ramsay and Silverman 2005](#));
- 100 log-periodograms of length 150 corresponding to recordings of speakers pronouncing the phoneme “aa” ([Ferraty and Vieu 2006](#));
- 76 nitrogen oxides (NO_x) emission level daily curves measured every hour close to an industrial area in Poblenou (Barcelona). All the curves correspond to working days ([Febrero et al. 2008](#)).

Before closing this introduction, we make a brief remark on a FDA aspect that is not central in this thesis, but it is worth to mention. According to [Cuevas \(2014\)](#), a functional datum $y(s)$ may require a preliminary treatment to either reduce its dimension or remove noise. Among other alternatives, both objectives can be achieved through basis representation: let $\{e_j(s)\}$ be a basis system of \mathbb{H} , then $y(s)$ can be substituted by a truncated basis representation given by

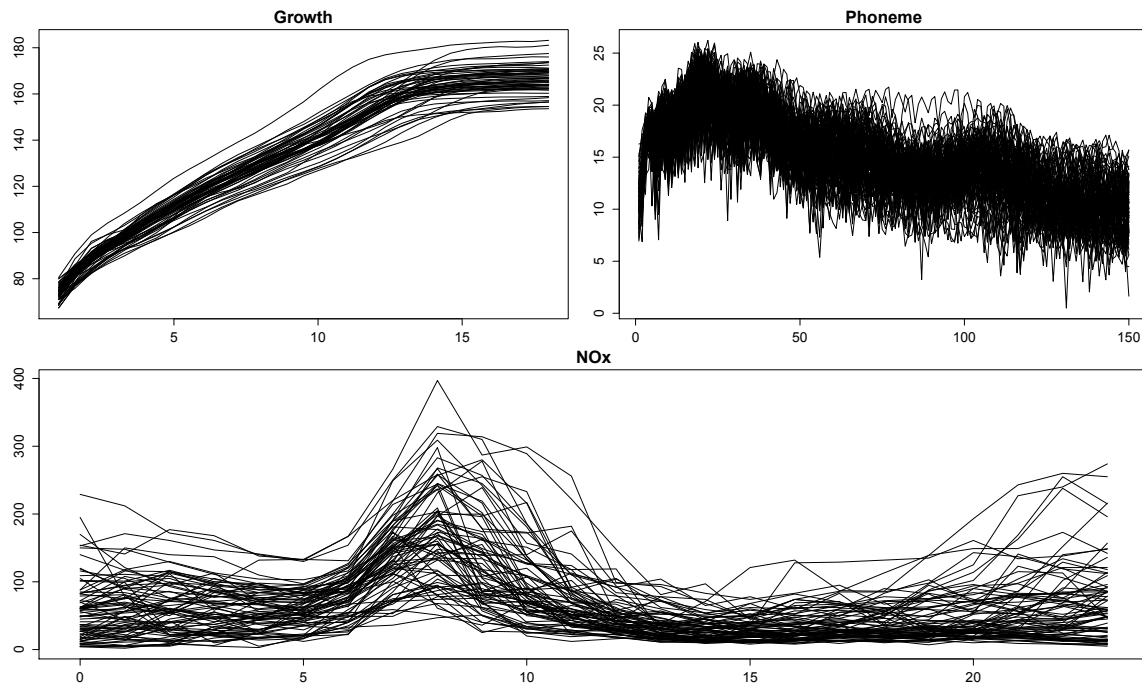


Figure 1.1: Three real functional datasets: growth curves of girls (top left), log-periodograms of the phoneme “aa” (top right), daily curves of NO_x levels (bottom)

$$\tilde{y}(s) = \sum_{j=1}^J c_j e_j(s),$$

where J denotes the number of basis functions to use and c_j the coefficients to be chosen according to some criteria, e.g., least squares

$$\sum_{k=1}^m \left(y(s_k) - \sum_{j=1}^J c_j e_j(s_k) \right)^2.$$

Clearly, the choices of $\{e_j(s)\}$, J and the criterion to select $\{c_j\}$ are very important aspects, but we do not discuss them in the thesis since they are beyond our scope. For more details on this FDA issue, we recommend the book by [Ramsay and Silverman \(2005\)](#).

1.1 The Notion of Functional Depth

In the previous section we discussed the juxtaposition of multivariate and functional data analysis. However, their differences did not prevent multivariate techniques to inspire advances in FDA. A good example of the collaboration between the two areas is given by the extension of the notion of data depth from the multivariate to the functional framework.

The idea of data depth was born in the multivariate context in a successful attempt to extend the univariate notion of order statistics to multidimensional spaces. Univariate order statistics allow to rank data from the smallest to the largest observation and to evaluate the degree of centrality of a point relative to a probability distribution or a sample. Moving to \mathbb{R}^d , $d \geq 2$, although \mathbb{R}^d lacks the natural order of \mathbb{R} , it is still interesting to have a center-outward ordering criterion and the notion of multivariate depth succeeds in providing it. According to [Serfling \(2006\)](#), a multivariate depth is a function that measures how deep (or central) a point $\mathbf{x} \in \mathbb{R}^d$ is relative to the probability distribution P . To give an idea of how a multivariate depth should work, consider a P with a unique mode \mathbf{M}_o and a certain tail. Then, the depth value of \mathbf{x} should be high for $\mathbf{x} = \mathbf{M}_o$ and low for \mathbf{x} equal to a value in the tail region.

Several implementations of the notion of multivariate depth have been proposed in the literature. For an overview on this topic, see for example [Liu et al. \(1999\)](#) or [Zuo and Serfling \(2000\)](#). Next, we report the definitions of three multivariate depths. First, the halfspace depth ([Tukey 1975](#)), which is defined as the minimum probability mass carried by any closed halfspace $H \in \mathbb{R}^d$ containing \mathbf{x} . More precisely,

Definition 1.1. Halfspace (or Tukey) depth ([Tukey 1975](#))

The halfspace depth of $\mathbf{x} \in \mathbb{R}^d$ relative to P on \mathbb{R}^d is given by

$$HSD(\mathbf{x}, P) = \inf_H \{\Pr(H) : \mathbf{x} \in H\}. \quad (1.1)$$

Second, the simplicial depth ([Liu 1990](#)), which is defined as the probability that \mathbf{x} belongs to a random simplex $S \in \mathbb{R}^d$, that is,

Definition 1.2. Simplicial depth (Liu 1990)

The simplicial depth of $\mathbf{x} \in \mathbb{R}^d$ relative to P on \mathbb{R}^d is given by

$$SID(\mathbf{x}, P) = \Pr(\mathbf{x} \in S(\mathbf{y}_1, \dots, \mathbf{y}_{d+1})), \quad (1.2)$$

where $S(\mathbf{y}_1, \dots, \mathbf{y}_{d+1})$ denotes the d -dimensional simplex with vertices $\mathbf{y}_1, \dots, \mathbf{y}_{d+1}$.

Third, the spatial depth (Serfling 2002), that uses the geometry of multivariate data clouds and is connected with a notion of multivariate quantiles as we show later.

Definition 1.3. Spatial depth (Serfling 2002)

Let \mathbf{Y} be a random variable having probability distribution P on \mathbb{R}^d and F be the associated cumulative distribution of \mathbf{Y} . The spatial depth of $\mathbf{x} \in \mathbb{R}^d$ relative to P is given by

$$SD(\mathbf{x}, P) = 1 - \left\| \int S(\mathbf{x} - \mathbf{y}) dF(\mathbf{y}) \right\|_E = 1 - \|\mathbb{E}[S(\mathbf{x} - \mathbf{Y})]\|_E, \quad (1.3)$$

where $\|\cdot\|_E$ is the Euclidean norm in \mathbb{R}^d and $S: \mathbb{R}^d \rightarrow \mathbb{R}^d$ is the spatial sign function given by

$$S(\mathbf{x}) = \begin{cases} \frac{\mathbf{x}}{\|\mathbf{x}\|_E}, & \mathbf{x} \neq \mathbf{0}, \\ \mathbf{0}, & \mathbf{x} = \mathbf{0}. \end{cases}$$

Among the previous three multivariate depths, SD has certainly the most important role for the development of this thesis. For this reason, we report two features of the multivariate spatial depth. First, as mentioned above, SD is connected with the notion of multivariate spatial quantile introduced by Chaudhuri (1996):

Definition 1.4. Spatial quantile (Chaudhuri 1996)

Let \mathbf{Y} be a random variable having probability distribution P on \mathbb{R}^d . Consider $\mathbf{u} \in \mathbb{R}^d$ such that $\|\mathbf{u}\|_E < 1$. Then, $Q_P(\mathbf{u})$ is the u th spatial quantile of \mathbf{Y} if and only if $Q_P(\mathbf{u})$ is the value of \mathbf{q} which minimizes

$$\mathbb{E}[\Psi(\mathbf{u}, \mathbf{Y} - \mathbf{q}) - \Psi(\mathbf{u}, \mathbf{Y})],$$

where, for $\mathbf{y} \in \mathbb{R}^d$, $\Psi(\mathbf{u}, \mathbf{y}) = \|\mathbf{y}\|_E + \langle \mathbf{u}, \mathbf{y} \rangle_E$, and $\langle \mathbf{u}, \mathbf{y} \rangle_E$ is the Euclidean inner product of \mathbf{u} and \mathbf{y} .

Then, it can be shown that under some mild conditions $Q_P(\mathbf{u})$ and $SD(\mathbf{x}, P)$ are linked in the following way:

$$\|Q_P^{-1}(\mathbf{x})\|_E = 1 - SD(\mathbf{x}, P). \quad (1.4)$$

Therefore, if \mathbf{x} has a high spatial depth value, its associated spatial quantile \mathbf{u} has a low norm, and vice versa.

Second, $SD(\mathbf{x}, P)$ depends on the whole P , that is, SD as well as HSD and SID tackle the depth problem through an approach that we describe as “global”. However, in some analyses it may be recommended to focus on narrower neighborhoods of P and tackle the depth problem with a “local approach”. With the aim of showing how a local approach can be implemented, we next briefly describe the local versions of the global depths HSD, SID and SD that have been proposed in the literature:

- the local halfspace depth LHSD (Agostinelli and Romanazzi 2011), which replaces closed halfspaces with closed slabs. Denote with $H_u(a)$ the closed halfspace $\{z \in \mathbb{R}^d : u'z \geq a, u'u = 1\}$. Then, (1.1) can be written as

$$HSD(\mathbf{x}, P) = \inf_{u:u'u=1} \Pr(H_u(u'x)). \quad (1.5)$$

A closed slab $SL_u(a, a + \tau)$ between two parallel closed halfspaces $H_u(a)$ and $H_u(a + \tau)$ is the intersection between the positive side of $H_u(a)$ and the negative side of $H_u(a + \tau)$.

Then, LHSD is given by the following modification of (1.5):

$$LHSD(\mathbf{x}, P, \tau) = \inf_{u:u'u=1} \Pr(SL_u(u'x, u'x + \tau)).$$

- the local simplicial depth LSID (Agostinelli and Romanazzi 2011), which only considers

simplices with size no greater than a fixed threshold τ and is given by the following modification of (1.2):

$$LSID(\mathbf{x}, P, \tau) = \Pr(\mathbf{x} \in S(\mathbf{y}_1, \dots, \mathbf{y}_{d+1}) : v(S(\mathbf{y}_1, \dots, \mathbf{y}_{d+1})) \leq \tau),$$

where $v(S(\mathbf{y}_1, \dots, \mathbf{y}_{d+1}))$ denotes the volume of the simplex with vertices $\mathbf{y}_1, \dots, \mathbf{y}_{d+1}$.

- the kernelized spatial depth KSD (Chen et al. 2009), which is basically based on considering $S(\mathbf{x} - \mathbf{y})$ in (1.3) through a kernel function that reduces the effect of \mathbf{y} on the depth of \mathbf{x} for \mathbf{y} distant from \mathbf{x} .

The notion of multivariate depth has been extended to functional data. In FDA, a depth has a similar purpose, that is, it allows to obtain a center-outward ordering criterion for curves, and different implementations already exist in the FDA literature. For example, Chakraborty and Chaudhuri (2014) defined the functional version of $SD(\mathbf{x}, P)$, the functional spatial depth function $FSD(x, P)$, where now x is an element of an infinite-dimensional Hilbert space \mathbb{H} and P is a probability distribution on \mathbb{H} . As well as SD, FSD is a global-oriented depth. We give its formal definition in Chapter 2, where we also present the functional extension of the notion of spatial quantile function, $FQ_P(u)$, where now $u \in \mathbb{H}$ such that $\|u\| < 1$, and $\|\cdot\|$ is the norm derived from the inner product $\langle \cdot, \cdot \rangle$ in \mathbb{H} (Chaudhuri 1996). In the same chapter we show that the relationship given by (1.4) extends to the functional framework, and therefore also $FSD(x, P)$ and $FQ_P(u)$ are linked.

The main contribution of this thesis is the definition of a new functional depth. Indeed, we define the kernelized functional spatial depth (KFSD) which can be interpreted in two ways: first, KFSD is a functional extension of the multivariate KSD proposed by Chen et al. (2009); second, KFSD is a local-oriented version of the global-oriented FSD. Then, we use KFSD to solve functional statistical problems that may require an analysis at a local level. We mainly focus on supervised classification and outlier detection and we show that KFSD represents a useful tool to solve such problems.

Along the thesis, besides FSD and KFSD, we consider other functional depths that we next

briefly introduce: [Fraiman and Muniz \(2001\)](#) defined the Fraiman and Muniz depth (FMD), which tries to measure how long a curve remains in the middle of a sample. [Cuevas et al. \(2006\)](#) proposed the h-modal depth (HMD), which tries to measure how densely a curve is surrounded by other curves. [Cuesta-Albertos and Nieto-Reyes \(2008\)](#) and [Cuevas and Fraiman \(2009\)](#) proposed the random Tukey depth (RTD) and the integrated dual depth (IDD), respectively. Both depths are based on the computation of K random one-dimensional projections of the curves, but they differ on how the projections are treated: RTD is given by the minimum of the univariate Tukey depth values of the projections, IDD is given by the average of the univariate simplicial depth values of the projections. Finally, [López-Pintado and Romo \(2009\)](#) proposed the band and modified band depths. In particular, the modified band depth (MBD) is based on all the possible bands defined by the graphs on the plane of $2, 3, \dots$ and J curves, and on a measure of the sets where another curve is inside these bands. We provide the definitions of these functional depths in [Chapter 2](#).

1.2 Depth-based Functional Methods

In what follows we often use the notion of i. i. d. functional sample that may refer alternatively to:

1. an i. i. d. random sample of size n generated from the random variable $Y \in \mathbb{H}$. In this case we use the notation $Y_n = \{y_1, \dots, y_n\}$;
2. an i. i. d. random sample of size n generated from the stochastic process $\{Y(s), s \in I\}$. In this case we use the notation $Y_n(s) = \{y_1(s), \dots, y_n(s)\}$;
3. an i. i. d. random sample of size n generated from the stochastic process $\{Y(s), s \in I\}$ whose i th component has been observed at $\mathbf{s}_i = (s_{1i}, \dots, s_{mi})$, $i \in \{1, \dots, n\}$. In this case we use the notation $Y_n(\mathbf{s}) = \{y_1(\mathbf{s}_1), \dots, y_n(\mathbf{s}_n)\}$.¹

¹if $\mathbf{s}_i \neq \mathbf{s}_j$ for some $i, j \in \{1, \dots, n\}$, it may be necessary a preliminary step to estimate the curves at a common set of domain points. In this thesis, when the data require such preliminary treatment, we carry out the estimation using cubic spline interpolation. Other techniques can be used for this task.

A functional sample (e.g., Y_n) can be analyzed using a functional depth. For example, functional location estimation of the center of Y can be provided by either non-depth-based or depth-based statistics. A natural estimation of the center of Y is given by the functional mean,

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n y_i.$$

Clearly, the computation of $\hat{\mu}$ does not require the use of a functional depth. However, a depth analysis of Y_n provides a center-outward ordering of the observations, i.e., $Y_{(n)} = \{y_{(1)}, \dots, y_{(n)}\}$, where $y_{(1)}$ and $y_{(n)}$ are the curves in Y_n with minimum and maximum depth, respectively. $Y_{(n)}$ provides an alternative to $\hat{\mu}$, that is, the functional depth-based median given by the deepest/most central observation in Y_n , $\hat{M}_e = y_{(n)}$. Another substitute for $\hat{\mu}$ may be the functional depth-based α -trimmed mean $\hat{\mu}_\alpha$, $0 < \alpha < 1$, that can be obtained as follows: compute $r_\alpha = \alpha n$ and take the smallest integer greater or equal than r_α , $\lceil r_\alpha \rceil$. Then,

$$\hat{\mu}_\alpha = \frac{1}{n} \sum_{i=\lceil r_\alpha \rceil}^n y_{(i)}.$$

Moreover, the information contained in $Y_{(n)}$ allows to estimate locations other than the center of Y by means of the notion of functional depth-based percentiles: let $0\% \leq p \leq 100\%$, $r_p = \frac{pn}{100}$ and \tilde{r}_p be the nearest integer to r_p , then the p depth-based percentile of Y_n is given by $y_{(\tilde{r}_p)}$.

Besides for location estimation, functional depths can be used to build methods that are robust to functional outliers, i.e., curves generated by a different random variable than the one of the normal curves. For instance, supervised functional classification is one of the problems that has been tackled using the notion of depth. In a supervised functional classification problem there are some labeled training curves that belong to two or more groups and the goal is to classify some test curves with unknown class membership. [López-Pintado and Romo \(2006\)](#) and [Cuevas et al. \(2007\)](#) proposed three depth-based classification methods which are especially devised to deal with functional samples where the presence of outlying curves cannot be discarded. They are the distance to the trimmed mean method (DTM, [López-Pintado and](#)

Romo 2006), the weighted averaged distance method (WAD, López-Pintado and Romo 2006) and the within maximum depth method (WMD, Cuevas et al. 2007). We formally present them in Chapter 3, but here we anticipate that the robustness to outliers of DTM, WAD and WMD is achieved thanks to the use of a functional depth: in DTM, the decision rule depends on the group depth-based trimmed means, that are resistant to outliers; in WAD, the classification is done giving more weight to those curves that are deeper within the group training samples, and outliers have usually low depth values; in WMD, the class of an unlabeled curve depends on its depth values relative to the different training groups, and these values are barely affected by outliers. We enhance the study of DTM, WAD and WMD by studying their performances when they are used in conjunction with KFSD and the other above-mentioned existing functional depths. In particular, we consider supervised functional classification problems where the differences between groups are not extremely clear-cut or the data may contain outlying curves, and we show that in these challenging scenarios a local approach based on the use of KFSD leads to good results.

An alternative strategy to the construction of robust functional methods is outlier detection. When a sample of curves is ordered from the most to the least central curve using a functional depth, if any outlier is in the sample, its depth is expected to be among the lowest values. Therefore, it is reasonable to build outlier detection methods based on this feature. Indeed, methods of this nature already exist in the literature. For example, Febrero et al. (2008) proposed to label as outliers those curves with depth values lower than a certain threshold. As functional depths, they considered three alternatives, i.e., the Fraiman and Muniz depth FMD (Fraiman and Muniz 2001), the h-modal depth HMD (Cuevas et al. 2006) and the integrated dual depth IDD (Cuevas and Fraiman 2009), and they proposed two bootstrap procedures based on depth-based trimmed and weighted resampling to determine the depth threshold. Also, Sun and Genton (2011) introduced the functional boxplot, which is constructed using the ordering provided by the modified band depth MBD (López-Pintado and Romo 2009). Then, the functional boxplot allows to detect outliers as well as the standard boxplot does. Clearly, the use of a functional depth is only one of the possible strategies for tackling the functional

outlier detection problem. For example, [Hyndman and Shang \(2010\)](#) proposed to reduce the outlier detection problem from functional to multivariate data by means of functional principal component analysis (FPCA), and to use two alternative multivariate techniques on the scores to detect outliers, i.e., the bagplot and the high density region boxplot. In this thesis, we contribute to the FDA literature on outlier detection by enlarging the number of available procedures. Indeed, we present three new methods that are based on smoothed resampling techniques and allow to select a threshold for KFSD to detect outliers. Similarly as for classification, we evaluate these new procedures in challenging scenarios: we focus on low magnitude, shape and partial outliers, that is, on outliers that are difficult to recognize, and we observe results that support the proposed KFSD-based outlier detection methods.

1.3 Structure of the Thesis

This thesis contains five chapters. In the current chapter we discussed some FDA issues, we presented three real functional datasets and we introduced the notion of functional depth. To do that, we also recalled some global and local multivariate depths and described how the concept of depth has been extended to FDA. Finally, we recalled some examples of how a functional depth can help in doing statistics in FDA, e.g., location estimation, supervised classification and detection of outliers.

The contributions of this dissertation are developed in Chapters [2](#), [3](#) and [4](#). In [Chapter 2](#) we present a new functional depth, the kernelized functional spatial depth (KFSD). KFSD is a depth relying on an approach that is both spatial and local. Its local orientation is based on the use of kernel methods that allow the depth value of a curve x to depend more on the curves lying in a narrow neighborhood of x . By means of motivating examples, we show that the local approach behind KFSD results useful to analyze functional samples having for instance a structure that deviates from unimodality or that are contaminated by outliers. In [Sections 2.3](#) and [2.2](#), we present the other existing functional depths that we use in this thesis (FMD, HMD, RTD, IDD, MBD and FSD).

In [Chapter 3](#) we use KFSD to perform supervised classification. After describing the statistical problem of functional discrimination, we present three existing depth-based methods (DTM, WAD and WMD) and carry out an extensive simulation study. The main feature of the study, and a novelty in the field, is that we focus on scenarios where the difference between groups are not extreme. We show that in these scenarios the use of KFSD together with some of the depth-based classification methods leads to competitive results. In [Section 3.4](#) we classify the curves of two real datasets and we confirm the good results observed with the simulated curves.

In [Chapter 4](#) we deal with outlier detection. As for classification, the strategy consists in using KFSD as a tool to reach our statistical goal. In this case the objective is the identification of outliers for which we provide three new different procedures based on KFSD. We define these methods in [Section 4.2](#), whereas we evaluate them in [Section 4.3](#), where we also consider some existing procedures as competitors. If in classification we do not consider classes extremely different, in outlier detection we not consider extreme outliers. Indeed, we find more interesting to focus on atypical curves that are difficult to recognize and the new procedures perform well in detecting this type of outliers. Finally, we close [Chapter 4](#) doing outlier detection on a real dataset.

The final chapter of the dissertation is dedicated to some conclusions and to the presentation of some possible future research lines.

È la vita, oggi a te domani a lui!
(Dante Cruciani to Ferribotte in *I Soliti Ignoti*)

Chapter 2

The Kernelized Functional Spatial Depth¹

2.1 Introduction

The origins of the general idea of spatial depth date back to [Brown \(1983\)](#), who studied the problem of robust location estimation for two-dimensional spatial data and introduced the idea of spatial median. The spatial approach considers the geometry of the data and is at the basis of the multivariate spatial depth ([Serfling 2002](#)) and quantiles ([Chaudhuri 1996](#)) introduced in [Chapter 1](#) (Definitions [1.3](#) and [1.4](#), respectively). Both notions have been extended to functional spaces and they are presented in [Section 2.5](#). In particular, we report the definition of the functional spatial depth (FSD) due to [Chakraborty and Chaudhuri \(2014\)](#). Then, we present a new functional depth, the kernelized functional spatial depth (KFSD), which arises from a modification of FSD. From now on, we assume that \mathbb{H} is an infinite-dimensional Hilbert space, therefore equipped with an inner product function $\langle \cdot, \cdot \rangle$ and a norm function $\| \cdot \|$ inherited from $\langle \cdot, \cdot \rangle$. Before presenting FSD and KFSD, we report the definitions of the other functional depths that we consider in this thesis.

¹This chapter is mostly based on the forthcoming article in the journal *TEST* by [Sguera et al. \(2014b\)](#) available online at <http://link.springer.com/article/10.1007/s11749-014-0379-1> or upon request (csguera@est-econ.uc3m.es)

2.2 Other Functional Depths

In this section we report the definitions of the other functional depths that we use to carry out depth-based supervised classification and outlier detection.

[Fraiman and Muniz \(2001\)](#) introduced the first implementation of the notion of depth for functional data and used it to define ranks and trimmed means in the functional framework. Their idea consists in considering the integral of the univariate depths of $x(s)$ at each single point $s \in I$:

Definition 2.1. Fraiman and Muniz depth (FMD, [Fraiman and Muniz 2001](#))

Let $\{Y(s), s \in I\}$ be a stochastic process in $C(I)$, the space of continuous functions on the interval I . The Fraiman and Muniz depth of $x(s) \in C(I)$ relative to $Y(s)$ is given by

$$FMD(x(s), Y(s)) = \int_I D(x(s), Y(s)) ds,$$

where $D(\cdot, \cdot)$ is a univariate depth. [Fraiman and Muniz \(2001\)](#) propose to use the univariate simplicial depth as $D(\cdot, \cdot)$, that is,

$$D(x(s), Y(s)) = F_s(x(s))(1 - F_s(x(s)))$$

for any $s \in I$, where $F_s(\cdot)$ is the cumulative distribution function of $Y(s)$ at any fixed $s \in I$.

In the initial attempt to extend the concept of mode to the functional setup, [Cuevas et al. \(2006\)](#) also defined a local depth based on the use of a kernel function. Their basic idea is to measure how densely a trajectory is surrounded by other trajectories of the process:

Definition 2.2. h-modal depth (HMD, [Cuevas et al. 2006](#))

Let Y be a functional random variable having probability distribution P on \mathbb{H} . The h-modal depth of $x \in \mathbb{H}$ relative to P is given by

$$HMD(x, P) = \mathbb{E}[\kappa_h(x, Y)] = \frac{1}{h} \mathbb{E}[\kappa(x, Y)],$$

where $\kappa_h, \kappa : \mathbb{H} \times \mathbb{H} \rightarrow \mathbb{R}$ are two kernel functions such that $\kappa_h(x, Y) = \frac{1}{h}\kappa(x, Y)$ and h is a fixed tuning parameter.

Clearly, HMD depends on the choice of κ and h . We report the recommendations of the authors in Chapters 3 and 4.

Cuesta-Albertos and Nieto-Reyes (2008) proposed a multivariate depth consisting in a random approximation of HSD based on a finite number of one-dimensional projections. Their first goal was to unburden the demanding computation of HSD, but they also defined a functional extension of their multivariate depth:

Definition 2.3. Random Tukey depth (RTD, Cuesta-Albertos and Nieto-Reyes 2008)

Let Y be a functional random variable having probability distribution P on \mathbb{H} . Let v be an absolutely continuous distribution on \mathbb{H} and $V = \{v_1, \dots, v_K\}$ be an i. i. d. random sample generated from v . The Random Tukey depth of $x \in \mathbb{H}$ relative to P on V is given by

$$RTD_V(x, P) = \min \{HSD(\langle v_k, x \rangle, P_{v_k}) : v_k \in V, k = 1, \dots, K\},$$

where P_{v_k} denotes the distribution of $\langle v_k, Y \rangle$.

Another random functional depth has been proposed by Cuevas and Fraiman (2009). We report its definition for Hilbert spaces, but this depth can also be defined for data that are elements of a Banach space.

Definition 2.4. Integrated dual depth (IDD, Cuevas and Fraiman 2009)

Let Y be a functional random variable having probability distribution P on \mathbb{H} . Let v be an absolutely continuous distribution on \mathbb{H} and $V = \{v_1, \dots, v_K\}$ be an i. i. d. random sample generated from v . The integrated dual depth of $x \in \mathbb{H}$ relative to P on V is given by

$$IDD_V(x, P) = \frac{1}{K} \sum_{k=1}^K SID(\langle v_k, x \rangle, P_{v_k}),$$

where P_{v_k} denotes the distribution of $\langle v_k, Y \rangle$.

Finally, [López-Pintado and Romo \(2009\)](#) introduced two definitions of depth for functional observations based on the graphic representation of the curves and on the bands in \mathbb{R}^2 delimited by j curves. First, the band depth:

Definition 2.5. Band depth (BD, [López-Pintado and Romo 2009](#))

Let $\{Y(s), s \in I\}$ be a stochastic process in $C(I)$ and $\{Y_1(s), \dots, Y_J(s)\}$ be J i. i. d. copies of $Y(s)$. The band depth of $x(s) \in C(I)$ relative to $Y(s)$ is given by

$$BD(x(s), Y(s)) = \sum_{j=2}^J \Pr \left(\min_{i=1, \dots, j} Y_i(s) \leq x(s) \leq \max_{i=1, \dots, j} Y_i(s), \forall s \in I \right),$$

In practice, BD consists in a sum of probabilities of indicator function values. Instead of considering the indicator function, [López-Pintado and Romo \(2009\)](#) introduced a more flexible definition by measuring the set where the function $x(s)$ is inside the corresponding band:

Definition 2.6. Modified band depth (MBD, [López-Pintado and Romo 2009](#))

Let $\{Y(s), s \in I\}$ be a stochastic process in $C(I)$ and $\{Y_1(s), \dots, Y_J(s)\}$ be J i. i. d. copies of $Y(s)$. The modified band depth of $x(s) \in C(I)$ relative to $Y(s)$ is given by

$$MBD(x(s), Y(s)) = \sum_{j=2}^J \mathbb{E} \left[\lambda \left(s \in I : \min_{i=1, \dots, j} Y_i(s) \leq x(s) \leq \max_{i=1, \dots, j} Y_i(s) \right) \right],$$

where $\lambda(\cdot)$ is the Lebesgue measure on I .

In this thesis we do not employ BD, but only MBD. As recommended by the authors, we use $J = 2$ as the maximum number of curves delimiting a band.

The last existing functional depth that we consider is the functional spatial depth. Since FSD is closely related to our proposal KFSD, we dedicate a whole section of the thesis to the introduction of FSD.

2.3 The Functional Spatial Depth and Associated Quantiles

The functional spatial depth FSD is a recent implementation of the general notion of functional depth:

Definition 2.7. Functional spatial depth (FSD, [Chakraborty and Chaudhuri 2014](#))

Let Y be a functional random variable having probability distribution P on \mathbb{H} . The functional spatial depth of $x \in \mathbb{H}$ relative to P is given by

$$FSD(x, P) = 1 - \|\mathbb{E}[FS(x - Y)]\|,$$

where $FS : \mathbb{H} \rightarrow \mathbb{H}$ is the functional spatial sign function given by²

$$FS(x) = \begin{cases} \frac{x}{\|x\|}, & x \neq 0, \\ 0, & x = 0. \end{cases}$$

Next, we show that FSD is connected with the notion of functional spatial quantiles due to [Chaudhuri \(1996\)](#):

Definition 2.8. Functional spatial quantile ([Chaudhuri 1996](#))

Let Y be a functional random variable having probability distribution P on \mathbb{H} . Consider $u \in \mathbb{H}$ such that $\|u\| < 1$. Then, $FQ_P(u)$ is the u th functional spatial quantile of Y if and only if $FQ_P(u)$ is the value of q which minimizes

$$\mathbb{E}[\Psi(u, Y - q) - \Psi(u, Y)], \quad (2.1)$$

where, for $y \in \mathbb{H}$, $\Psi(u, y) = \|y\| + \langle u, y \rangle$.

[Cardot et al. \(2013\)](#) showed that, if Y is not concentrated on a straight line and is not strongly concentrated around single points, the Fréchet derivative of the convex function (2.1) is given by

$$\psi(q) = -\mathbb{E} \left[\frac{Y - q}{\|Y - q\|} \right] - u. \quad (2.2)$$

Therefore, $FQ_P(u)$ is also the value of q such that $\psi(q) = 0$. Let x be the solution of $\psi(q) = 0$. Then, $-\mathbb{E} \left[\frac{Y - x}{\|Y - x\|} \right] = u = FQ_P^{-1}(x)$ and

²For $x = x(s) = 0$ for any $s > 0$ we use the notation $x = 0$.

$$\|FQ_P^{-1}(x)\| = \left\| -\mathbb{E} \left[\frac{Y - x}{\|Y - x\|} \right] \right\| = \|\mathbb{E}[FS(x - Y)]\| = 1 - FSD(x, P),$$

which indeed shows that the direct connection between the notions of spatial depth and quantiles holds also in functional Hilbert spaces. Besides this interesting interpretability property of FSD, [Chakraborty and Chaudhuri \(2014\)](#) also showed that: (1) $FSD(x, P)$ is invariant under the class of linear transformations $T : \mathbb{H} \rightarrow \mathbb{H}$, where $T(x) = cAx + b$, for $c \in \mathbb{R}$, $c > 0$, $b \in \mathbb{H}$ and an isometry A on \mathbb{H} ; (2) if P is non-atomic, then $FSD(x, P)$ is continuous in x ; (3) if \mathbb{H} is strictly convex and P is non-atomic and not supported on a line in \mathbb{H} , then $FSD(x, P)$ has a unique maximum at the spatial median M_e of Y and its maximum value is 1;³ (4) for any non-zero $x \in \mathbb{H}$ and sequence $\{M_e + nx\}_{n \in \mathbb{N}^+}$, the following holds: $FSD(m + nx, P) \rightarrow 0$ as $n \rightarrow \infty$; (5) $FSD(x, P)$ does not suffer from degeneracy for many infinite dimensional probabilities distributions. For more details on the desirable properties of a functional depth, see [Mosler and Polyakova \(2012\)](#).

When a functional sample is observed, $FSD(x, P)$ is replaced by its corresponding sample version:

Definition 2.9. Functional spatial depth, sample version ([Chakraborty and Chaudhuri 2014](#))

Let $Y_n = \{y_1, \dots, y_n\}$ be an i. i. d. random sample generated from the random variable $Y \in \mathbb{H}$.

The sample functional spatial depth of $x \in \mathbb{H}$ relative to Y_n is given by

$$FSD(x, Y_n) = 1 - \frac{1}{n} \left\| \sum_{i=1}^n FS(x - y_i) \right\|. \quad (2.3)$$

Note that the expression of above will result important for the definition of KFSD.

2.4 The Kernelized Functional Spatial Depth

In this section we propose the kernelized functional spatial depth, a local-oriented version of FSD. A common way to implement a local approach is to consider kernel-based methods. This

³If Y is not concentrated on a straight line and is not strongly concentrated around single points, the spatial median M_e of Y is the unique solution of (2.2) for $u = 0$.

was the strategy of [Chen et al. \(2009\)](#), who proposed the multivariate kernelized spatial depth KSD, the local version of the multivariate spatial depth SD. Moving to functional spaces, we achieve a similar result and we provide a local depth based on FSD.

To do this, our first step consists in recoding the data. More in detail, instead of considering $x \in \mathbb{H}$, we consider $\phi(x) \in \mathbb{F}$, where $\phi : \mathbb{H} \rightarrow \mathbb{F}$ is an embedding map and \mathbb{F} is a feature space. Note that ϕ can be defined implicitly by a positive definite and stationary kernel, $\kappa : \mathbb{H} \times \mathbb{H} \rightarrow \mathbb{R}$, through

$$\kappa(x, y) = \langle \phi(x), \phi(y) \rangle.$$

For this reason, we refer to this new functional depth as kernelized functional spatial depth. Its definition is the following:

Definition 2.10. Kernelized functional spatial depth

Let Y be a functional random variable having probability distribution P on \mathbb{H} , $\phi : \mathbb{H} \rightarrow \mathbb{F}$ be an embedding map and \mathbb{F} be a feature space. The kernelized functional spatial depth of $x \in \mathbb{H}$ relative to P is given by

$$KFSD(x, P) = 1 - \|\mathbb{E}[FS(\phi(x) - \phi(Y))]\| = FSD(\phi(x), P_\phi),$$

where $\phi(Y)$ and P_ϕ are recoded version of Y and P , respectively.

The sample version of $KFSD(x, P)$ is easily obtainable and given by

$$KFSD(x, Y_n) = 1 - \frac{1}{n} \left\| \sum_{i=1}^n FS(\phi(x) - \phi(y_i)) \right\| = FSD(\phi(x), \phi(Y_n)), \quad (2.4)$$

where $\phi(Y_n)$ is a recoded version of Y_n . Therefore, both $KFSD(x, P)$ and $KFSD(x, Y_n)$ can be interpreted as recoded versions of $FSD(x, P)$ and $FSD(x, Y_n)$, respectively.

Note that for $KFSD(x, Y_n)$ it is possible to obtain an expression where its kernel nature is explicit. Indeed, the following holds:

$$\left\| \sum_{i=1}^n FS(x - y_i) \right\|^2 = \sum_{\substack{i,j=1; \\ y_i \neq x; y_j \neq x}}^n \frac{\langle x, x \rangle + \langle y_i, y_j \rangle - \langle x, y_i \rangle - \langle x, y_j \rangle}{\sqrt{\langle x, x \rangle + \langle y_i, y_i \rangle - 2\langle x, y_i \rangle} \sqrt{\langle x, x \rangle + \langle y_j, y_j \rangle - 2\langle x, y_j \rangle}},$$

which implies that $FSD(x, Y_n)$ can be expressed in terms of inner products. Thereby, recoding the data is equivalent to consider a kernel instead of the inner product. Both inner products and kernel functions can be seen as similarity measures, but kernels are more powerful and richer than inner products. Therefore, we opt for replacing the inner product function with a positive definite and stationary kernel function, thus obtaining a κ -based sample version of KFSD:

Definition 2.11. Kernelized functional spatial depth, κ -based sample version

Let $Y_n = \{y_1, \dots, y_n\}$ be an i. i. d. random sample generated from the random variable $Y \in \mathbb{H}$ and κ be positive definite and stationary kernel, $\kappa : \mathbb{H} \times \mathbb{H} \rightarrow \mathbb{R}$. The κ -based sample kernelized functional spatial depth of $x \in \mathbb{H}$ relative to Y_n is given by

$$KFSD(x, Y_n) = 1 - \frac{1}{n} \left(\sum_{\substack{i,j=1; \\ y_i \neq x; y_j \neq x}}^n \frac{\kappa(x, x) + \kappa(y_i, y_j) - \kappa(x, y_i) - \kappa(x, y_j)}{\sqrt{\kappa(x, x) + \kappa(y_i, y_i) - 2\kappa(x, y_i)} \sqrt{\kappa(x, x) + \kappa(y_j, y_j) - 2\kappa(x, y_j)}} \right)^{1/2}. \quad (2.5)$$

Note that (2.5) only requires the choice of κ , and not of ϕ , which can be left implicit. For this reason, in practice we use the κ -based version of $KFSD(x, Y_n)$.

Regarding the choice of κ , we use a functional version of the Gaussian kernel function used by [Chen et al. \(2009\)](#), that is,

$$\kappa(x, y) = \exp\left(-\frac{\|x - y\|^2}{\sigma^2}\right),$$

which in turn depends on the norm inherited by the functional Hilbert space where data are assumed to lie and on the bandwidth σ . With the aim of exploring the behavior of κ as a function of bandwidth, we consider different values of σ . We set each σ equal to the $p\%$ percentile of the empirical distribution of $\{\|y_i - y_j\|, y_i, y_j \in Y_n\}$. Note that the lower p , the lower σ and the more local the approach of KFSD. Indeed, with low percentiles KFSD considers small neighborhoods and can be viewed as a potential functional density estimator, whereas with high percentiles KFSD considers wide neighborhoods and behaves almost as a global depth. Therefore, the use of different percentiles allows us to cover different degrees of KFSD-based local approaches. For example, the 10%, 50% and 90% percentiles lead to strongly, moderately and weakly local approaches, respectively. In the corresponding chapters, we present two methods to choose a final value of σ in supervised classification and outlier detection problems.

After recalling the definitions of FMD, HMD, RTD, IDD, MBD and FSD, and after presenting KFSD, we dedicate the next section to discuss the differences between global and local depths.

2.5 From Global to Local Depths

The definition of $FSD(x, Y_n)$ given by (2.3) allow us to discuss an important feature of FSD, that is, its global approach to the depth problem, and to show how a local approach is an alternative. The functional spatial depth of x relative to Y_n depends on the values of the functional spatial sign function, $FS(x - y_i), i = 1, \dots, n$. Each $FS(x - y_i)$ is a unit-norm curve representing the direction from x to y . Therefore, $FSD(x, Y_n)$ depends equally on n directions. This feature generates a trade-off: on one side, $FSD(x, Y_n)$ turns out to be robust to the presence of outliers in Y_n ; on the other side, $FSD(x, Y_n)$ transforms $(x - y_i)$ into a unit-norm curve regardless of y_i being a neighboring or a distant curve from x . For this reason, we describe FSD as a global-oriented depth since it makes depend the depth of x on the whole Y_n , and with equal weights. However, in some circumstances it may be useful an analysis of narrow neighborhoods of x and to give more weight to close than distant curves, in other words, to have a local approach

to the depth problem that allow the information brought by y_i to depend on the value of a certain distance between x and y_i , as it happens with KFSD.

To illustrate the differences between a global and a local approach to the depth problem, we show two examples where the goal is to obtain a center-outward ordering of functional data. We consider five global depths (FMD, RTD, IDD, MBD and FSD) and two local depths (HMD and our proposal KFSD). In the first example, we generated 21 curves from a given process and divided them in three groups with 10, 10 and 1 curves, respectively. Then, we added a different constant curve to each group (the constants are 0, 10 and 5, respectively), obtaining the curves at the top of [Figure 2.1](#). Afterwards, we computed the depth values of all the curves using the five global-oriented depths and we observed that according to all the global depths the highest depth value is attained at the curve of the third group.

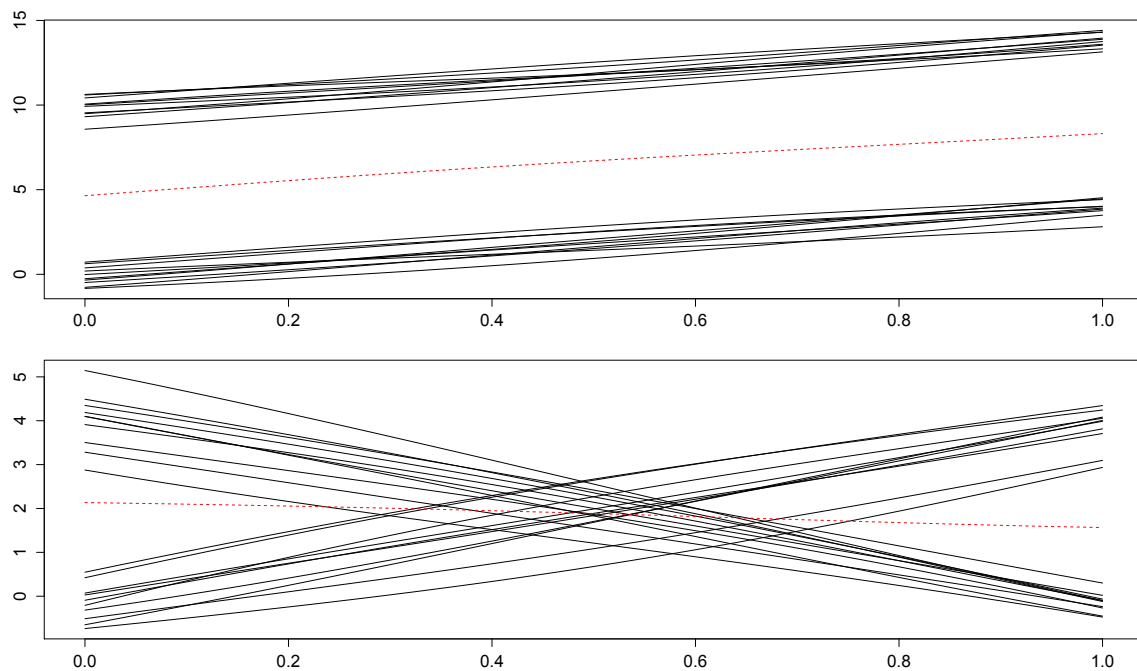


Figure 2.1: Two functional datasets that we use to show the differences between global and local depths

The structure of the dataset of the second example is similar, but we used different transformations to obtain the curves belonging to the second and third group (see the plot at the bottom of [Figure 2.1](#)). Also in this case, we observed that according to all global depths the

curve of the third group turns out to be the deepest one. Clearly, both examples involve three strongly different classes of curves. Nevertheless, when we treat the curves of each example as belonging to a homogeneous sample, we observe rather inconvenient behaviors of the global depths. In the two examples, the curve in the third group is roughly in the geometric center of the dataset, but it is far from the remaining curves and it would be more reasonable to observe a low depth value at this curve. Actually, this happens with the local depths HMD and KFSD, and mainly because these depths reduce the contribution of distant curves to the depth value of a given curve. Indeed, in both examples HMD and KFSD assigned the lowest depth value to the curves belonging to the third group. This peculiar behavior of HMD and KFSD is due to their local approach to the depth problem.

Additionally, we also want to give an idea of the potential usefulness of local depths in ranking correctly outliers that contaminate a functional sample, but that are somehow graphically hidden. To illustrate this fact, we present the following examples: first, we generated 10 datasets of size 50 from a mixture of two stochastic processes, one for normal curves and one for outliers that are graphically clear, with the probability that a curve is an outlier equal to 0.05. Second, we generated another group of 10 datasets from a different mixture which produces outliers that are instead graphically faint. In Figure 2.2 we report a contaminated dataset for each mixture.

Let $n_{out,j}, j = 1, \dots, 10$, be the number of outliers generated in the j th dataset. For each dataset and functional depth, it is desirable to assign the $n_{out,j}$ lowest depth values to the $n_{out,j}$ generated outliers. For both mixtures and each generated dataset, we registered how many times the depth of an outlier is indeed among the $n_{out,j}$ lowest values. As depth functions, we considered the five global depths and two local depths mentioned above. The results reported in Table 2.1 show that for all the functional depths the ranking of clear outliers is an easier task than the ranking of faint outliers.

However, while the ranking of clear outliers is reasonably good in different cases, e.g., local KFSD (100%) or global RTD (95%), the ranking of faint outliers is markedly better with local depths, i.e., HMD and KFSD (both 76.19%). These results give an idea of the potential of local

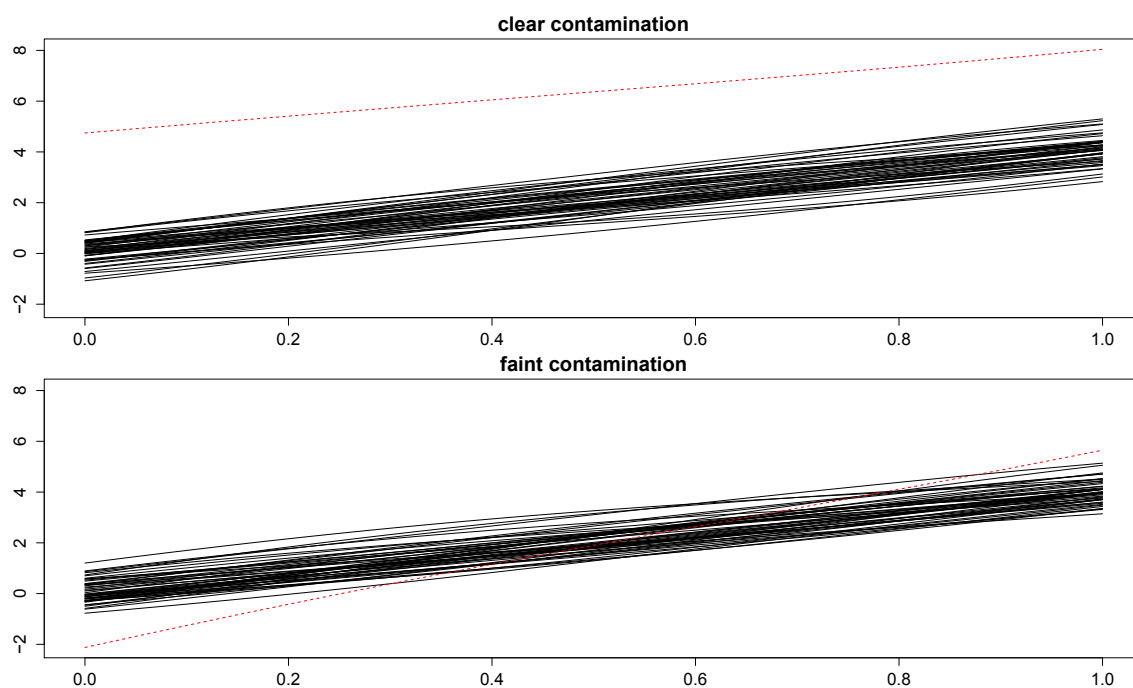


Figure 2.2: Examples of contaminated datasets: clear contamination (top) and faint contamination (bottom). The solid curves are normal curves and the dashed curves are outliers

Table 2.1: Percentages of times a depth assigns a value among the $n_{out,j}$ lowest ones to an outlier. Types of outliers: clear and faint

type of depths	global depths					local depths	
	FMD	RTD	IDD	MBD	FSD	HMD	KFSD
clear outliers	85.00	95.00	70.00	60.00	60.00	85.00	100.00
faint outliers	0.00	28.57	38.10	14.29	33.33	76.19	76.19

depths in detecting correctly faint outliers.

Finally, we present a comparative example to evaluate the rankings associated to KFSD and the other functional depths (FMD, HMD, RTD, IDD, MBD and FSD). We used the 54 growth curves already reported in [Figure 1.1](#) as functional sample Y_n . We computed their depth values using KFSD with a strongly local percentile (i.e., 10%), as well as using the remaining functional depths. Then, for each depth we obtained $Y_{(n)}$, that is, the depth-based ordered version of Y_n , and therefore the ranks of the curves (the curve with rank equal to 1 is the deepest curve, and so forth). We compared the different rankings and we summarized these comparisons in [Figure 2.3](#). For example, the figure at the top on the left of [Figure 2.3](#) compares the KFSD-based ranks (horizontal axis, in an increasing order) with the FMD-based ranks (in the vertical axis).

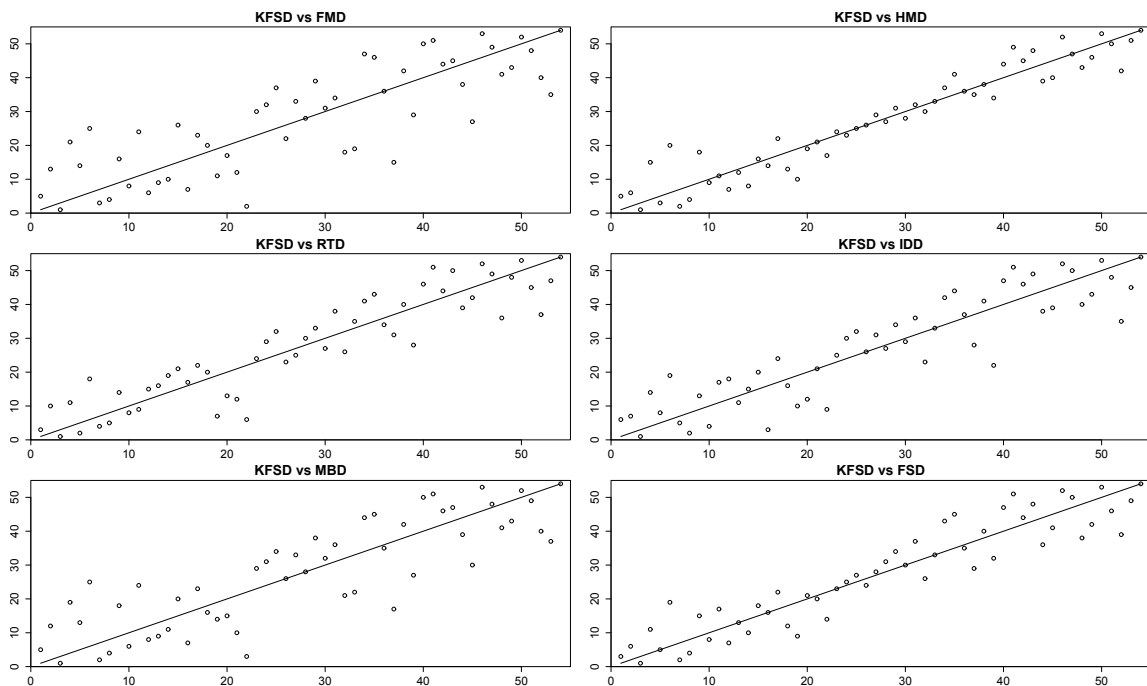


Figure 2.3: Comparison of depth-based rankings (KFSD versus FMD, HMD, RTD, IDD, MBD and FSD) using the 54 growth curves of [Figure 1.1](#)

Observing [Figure 2.3](#), it is possible to make some remarks:

- The KFSD-based ranking is rather different from the other rankings. However, it seems that HMD and FSD are the closest depth in terms of ranking similarities. This result is

not surprising since HMD is also a local depth, whereas FSD is the global counterpart of KFSD.

- The KFSD-based median, that is, the curve with rank equal to 1 according to KFSD, is not the median for any other depth. On the contrary, all the depths agree on which is the least deep curve, that is, the curve with rank equal to 54.

We also evaluated different KFSD-based local approaches, As mentioned above, to set σ in KFSD we consider different percentiles of $\{\|y_i - y_j\|, y_i, y_j \in Y_n\}$. For example, looking at the 10%, 50% and 90% percentiles it is possible to evaluate if strongly, moderately and weakly KFSD-based local approaches differ among them. Using the same functional dataset, we did this comparison, which is reported in Figure 2.4. The figure at the top of Figure 2.4 compares the KFSD-based ranks using the 10% percentile (horizontal axis, in an increasing order) with the case of the 50% percentile (in the vertical axis), while at the bottom the comparison is with the case of the 90% percentile.

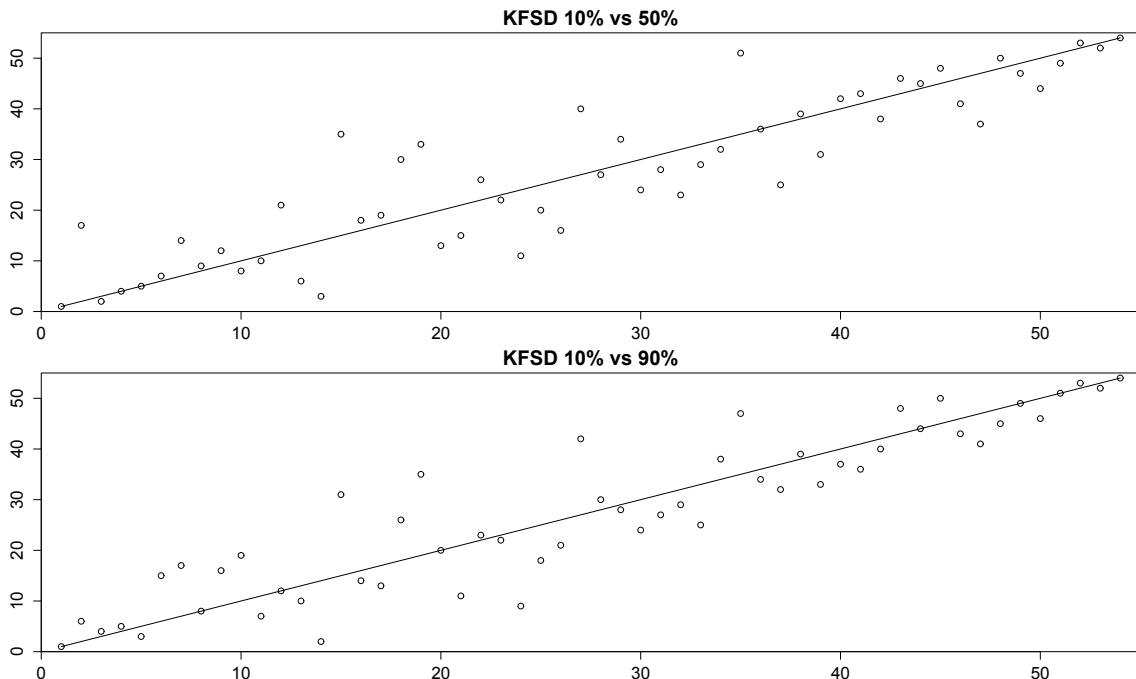


Figure 2.4: Comparison of KFSD-based rankings (percentile 10% versus 50% and 90%) using the 54 growth curves of Figure 1.1

Observing [Figure 2.4](#), it is possible to appreciate that the different percentiles generate different rankings. Therefore, the way how to choose an appropriate percentile for KFSD is an issue to take into account, as indeed we do in the rest of the thesis.

2.6 Conclusions

In this chapter we introduced the kernelized functional spatial depth, a local-oriented and kernel-based version of the functional spatial depth recently proposed by [Chakraborty and Chaudhuri \(2014\)](#). Originally developed by [Chaudhuri \(1996\)](#) and [Serfling \(2002\)](#) in the multivariate context, the spatial approach allows both FSD and KFSD to study the degree of centrality of curves from a new point of view with respect to the other existing functional depths. The main novelty introduced by KFSD consists in the fact that it addresses the study of functional datasets at a local spatial level, whereas FSD is more appropriate for global spatial analyses. For example, by means of two illustrative examples, we showed the potential of KFSD in analyzing functional samples that deviate from unimodality or in ranking correctly outliers graphically hidden.

As we showed in [Section 2.4](#), KFSD and FSD are related: $KFSD(x, P) = FSD(\phi(x), P_\phi)$, where $\phi : \mathbb{H} \rightarrow \mathbb{F}$ is an embedding map, \mathbb{F} is a feature space and P_ϕ is a recoded version of the probability distribution P . The embedding map ϕ and the feature space \mathbb{F} are implicitly defined through a positive definite kernel, $\kappa : \mathbb{H} \times \mathbb{H} \rightarrow \mathbb{R}$. This relationship between FSD and KFSD may represent the basis for the future theoretical study of the properties of KFSD.

KFSD depends on the choice of κ . Another possible future research line consists in the analysis of some alternatives to the kernel function employed in this thesis. Regarding the κ that we use at the moment, it depends on its bandwidth σ . To understand the effects of σ on the behavior of KFSD, we proposed to consider a range of values for σ . This choice allows to cover different degrees of KFSD-based local approaches, from strongly to weakly local approaches, and we showed that they may provide different depth rankings. The role and the choice of σ in classification and outlier detection will be further analyzed in [Chapters 3 and 4](#).

*Our nada who art in nada,
Nada be thy name.
Thy kingdom nada.
Thy will be nada
in nada as it is in nada.
(Ernest Hemingway, A Clean Well Lighted Place)*

Chapter 3

Supervised Functional Classification¹

3.1 Introduction

In this chapter we use KFSD as a tool to perform supervised functional classification. In general, in a supervised functional classification problem there is available a training sample composed of curves belonging to two or more groups and with known class memberships. The information contained in the training data is then used to assign test curves with unknown class membership to one of the groups. The analysis of the training information and the construction of the classification rule may be based on the use of a functional depth. Indeed, in this chapter we consider three existing depth-based procedures to perform supervised functional classification with simulated and real data. We apply these methods using a battery of functional depths: KFSD, but also FMD, HMD, RTD, IDD, MBD and FSD. We focus on challenging scenarios in which the differences between groups are not extremely clear-cut or the data may contain outlying curves. The results that we observe with simulated and real data indicate that a local KFSD-based approach is a good solution in such setups.

¹This chapter is mostly based on the forthcoming article in the journal *TEST* by [Sguera et al. \(2014b\)](#) available online at <http://link.springer.com/article/10.1007/s11749-014-0379-1> or upon request (csguera@est-econ.uc3m.es)

3.2 The Supervised Functional Classification Problem

The natural theoretical framework of supervised functional classification is given by the random pair (Y, G) , where Y is a functional random variable and G is a categorical random variable describing the class membership. From now on, we assume that G takes values 0 or 1. We also assume to observe a sample of n independent pairs generated from (Y, G) , i.e., $(Y_n, G_n) = \{(y_1, g_1), \dots, (y_n, g_n)\}$, with n_0 observations from the group with label 0 and n_1 observations from the group with label 1, $n_0 + n_1 = n$, and a curve x with unknown class membership generated from Y .

Using the information contained in (Y_n, G_n) , the goal of any supervised functional classification method is to provide a rule to classify the curve x , and several methods have been proposed in the literature. For instance, [Hastie et al. \(1995\)](#) have proposed a penalized version of the multivariate linear discriminant analysis technique, whereas [James and Hastie \(2001\)](#) have directly built a functional linear discriminant analysis procedure that uses natural cubic spline functions to model the observations. Using a P-spline approach, [Marx and Eilers \(1999\)](#) have considered functional supervised classification as a special case of a generalized linear regression model. [Hall et al. \(2001\)](#) have suggested to perform dimension reduction by means of functional principal component analysis and then to solve the derived multivariate problem with quadratic discriminant analysis or kernel methods. [Ferraty and Vieu \(2003\)](#) have developed a functional kernel-type classifier. [Epifanio \(2008\)](#) have proposed to describe curves by means of shape feature vectors and to use classical multivariate classifiers for the discrimination stage. [Galeano et al. \(2014\)](#) have developed new versions of several well known functional classification procedures which use a semi-distance for functional observations that generalizes the multivariate Mahalanobis distance. Finally, [Biau et al. \(2005\)](#) and [Cérou and Guyader \(2006\)](#) have studied some consistency properties of the extension of the k -nearest neighbor procedure to infinite-dimensional spaces. The first extension considers a reduction of the dimensionality of the regressors based on a Fourier basis system, and the second generalization deals with the real infinite dimension of the spaces under consideration. Note that the k -nearest neighbor method is indeed a general tool that can be used to perform functional nonparamet-

ric regression, and that classification is a specific case that occurs when the response of the regression model is categorical (for more details and theoretical results on the general case, see [Burba et al. 2009](#) and [Kudraszow and Vieu 2013](#)).

Besides the above described methods, other alternatives specially designed for datasets that may contain outlying curves and based on the use of a functional depth have been proposed. Indeed, in [Section 3.3](#) we carry out a simulation study with scenarios that allow outliers, whereas in [Section 3.4](#) we consider two potentially contaminated real datasets. Next, we report three depth-based proposals:

Definition 3.1. Distance to the trimmed mean method (DTM, [López-Pintado and Romo 2006](#))

Let $(Y_n, G_n) = \{(y_1, g_1), \dots, (y_n, g_n)\}$ be an i. i. d. random sample generated from the random pair $(Y, G) \in \mathbb{H} \times \{0, 1\}$ and x a random observation generated from the random variable $Y \in \mathbb{H}$. The distance to the trimmed mean method works as follows:

1. compute the depth-based α -trimmed means $\hat{\mu}_{\alpha, g}, g = 0, 1$;
2. compute $d_g = \|x - \hat{\mu}_{\alpha, g}\|, g = 0, 1$;
3. assign x to the group minimizing d_g .

Definition 3.2. Weighted averaged distance method (WAD, [López-Pintado and Romo 2006](#))

Let $(Y_n, G_n) = \{(y_1, g_1), \dots, (y_n, g_n)\}$ be an i. i. d. random sample generated from the random pair $(Y, G) \in \mathbb{H} \times \{0, 1\}$ and x a random observation generated from the random variable $Y \in \mathbb{H}$. Denote with $Y_{n_g} = \{y_{1_g}, \dots, y_{n_g}\}$ the functional sample composed of curves belonging to group $g, g = 0, 1$. The weighted averaged distance method works as follows:

1. for a given functional depth, say $FD(\cdot, \cdot)$, compute the within-group depth values, that is, $FD(y_{i_g}, Y_{n_g}), i = 1, \dots, n_g, g = 0, 1$. Let $w_{i_g} = FD(y_{i_g}, Y_{n_g})$;
2. compute $W_g = \sum_{i_g=1}^{n_g} \frac{w_{i_g} \|x - y_{i_g}\|}{\sum_{i_g=1}^{n_g} w_{i_g}}, g = 0, 1$;
3. assign x to the group minimizing W_g .

Definition 3.3. Within maximum depth method (WMD, Cuevas et al. 2007)

Let $(Y_n, G_n) = \{(y_1, g_1), \dots, (y_n, g_n)\}$ be an i. i. d. random sample generated from the random pair $(Y, G) \in \mathbb{H} \times \{0, 1\}$ and x a random observation generated from the random variable $Y \in \mathbb{H}$. The within maximum depth method works as follows:

1. for a given functional depth, compute the depth values of x relative to Y_{n_g} , that is, $FD(x, Y_{n_g}), g = 0, 1$. Let $D_g = FD(x, Y_{n_g})$;
2. assign x to the group maximizing D_g .

DTM, WAD and WMD can be used together with any functional depth. In the next two sections we compare the performances of these methods when used in conjunction with alternative functional depths such as FMD, HMD, RTD, IDD, MBD, FSD and KFSD. Among the non-depth-based methods, we consider as benchmark the functional k -nearest neighbor procedure (k -NN, Cérou and Guyader 2006), which looks at the k nearest neighbors of x in Y_n in terms of the norm of \mathbb{H} and assigns x to a group according to the majority vote. It is worth noting that the classification rule characterizing k -NN makes the method rather robust to the presence of outliers, and this is the reason why k -NN represents an interesting competitor for DTM, WAD and WMD.

3.3 Simulation Study

In this chapter and Chapter 2 we have presented three different depth-based classification procedures (DTM, WAD and WMD) and seven different functional depths (FMD, HMD, RTD, IDD, MBD, FSD and KFSD). Pairing all the procedures with all the depths, we obtain 21 depth-based classification methods, plus k -NN. The goal of this section is to compare them through an extensive simulation study. From now on, we refer to each method by the notation procedure+depth: for example, DTM+FMD refers to the method obtained by using DTM together with FMD.

We mainly explore the effectiveness of the methods in supervised classification problems

where the appropriate class membership of the curves is hard to be deduced by using graphical tools and/or in scenarios in which outlying curves are allowed. In these cases, it may happen that the curve we want to classify is geometrically rather central relative to the training samples of two different groups, but also relatively far from one of them, although not in an obvious way. In such scenarios, a depth-based classification method may behave better when used together with a local depth instead of a global depth.

Both methods and depths may depend on some parameters or assumptions. Regarding the methods, DTM depends on the trimming parameter α , that we set at $\alpha = 0.2$, as in [López-Pintado and Romo \(2006\)](#). For the benchmark procedure k -NN, we take $k = 5$ nearest neighbors since it is a standard choice and the method is reasonably robust with respect to the parameter k . Regarding the functional depths, for HMD, we follow the recommendations in [Febrero et al. \(2008\)](#), that is, \mathbb{H} is the L^2 space, $\kappa(x, y) = (2/\sqrt{2\pi}) \times \exp(-\|x - y\|^2/2h^2)$, and h is equal to the 15% percentile of the empirical distribution of $\{\|y_i - y_j\|, i, j = 1, \dots, n\}$. Note that for WMD+HMD we use a normalized version of HMD to make its range equal to $[0, 1]$. For RTD and IDD, we work with 50 projections in random Gaussian directions. For MBD, we consider bands defined by two curves. For FSD and KFSD, we assume that the curves lie in the L^2 space. Moreover, as introduced in [Chapter 2](#), we consider a set of percentiles $\{p_k\} = \{p_1, \dots, p_K\}$ to set σ in KFSD. In particular, we take $K = 7$ and $\{p_k\} = \{15\%, 25\%, 33\%, 50\%, 66\%, 75\%, 85\%\}$.

Next, we describe the structure of our simulation study. We overlook setups in which curves may be almost well classified by a preliminary graphical analysis and focus on classification scenarios where the differences among groups are hard to be detected graphically. Moreover, we allow outliers in one part of the simulation study. We consider two-groups scenarios throughout the whole simulation study, that is, $g \in \{0, 1\}$, and $x_g(s)$ denotes the curve generating process for group g .

In absence of contamination, we initially consider two different pairs of curve generating processes:

1. First pair of curve generating processes (from now on, CGP1) with $s \in [0, 1]$

$$\begin{aligned}x_0(s) &= 4s + \epsilon(s), \\x_1(s) &= 8s - 2 + \epsilon(s),\end{aligned}$$

where $\epsilon(s)$ is a zero-mean Gaussian component with covariance function given by

$$\mathbb{E}(\epsilon(s), \epsilon(s')) = 0.25 \exp(-(s - s')^2), \quad s, s' \in [0, 1]. \quad (3.1)$$

2. Second pair of curve generating processes (from now on, CGP2) with $s \in [0, 2\pi]$

$$\begin{aligned}x_0(s) &= u_{01} \sin s + u_{02} \cos s, \\x_1(s) &= u_{11} \sin s + u_{12} \cos s,\end{aligned} \quad (3.2)$$

where u_{01} and u_{02} are observations from a continuous uniform random variable between 0.05 and 0.1, whereas u_{11} and u_{12} are observations from a continuous uniform random variable between 0.1 and 0.12.

The main difference between the two processes is that for CGP1 both $x_0(s)$ and $x_1(s)$ are composed of deterministic and linear mean functions, plus a random component, while for CGP2 both $x_0(s)$ and $x_1(s)$ are exclusively composed of random and nonlinear mean functions.

To allow contamination, we consider the following modified version of CGP1 and CGP2, where the contamination affects only group 0:

1. First pair of curve generating processes allowing outliers (from now on, CGP1_{out}) with $s \in [0, 1]$

$$\begin{aligned}x_0(s) &= \begin{cases} 4s + \epsilon(s), & \text{with probability } 1 - q, \\ 4\sqrt{s} + \epsilon(s), & \text{with probability } q. \end{cases} \\x_1(s) &= 8s - 2 + \epsilon(s),\end{aligned}$$

where $\epsilon(s)$ is a zero-mean Gaussian component with covariance function given by (3.1) and $0 < q < 1$.

2. Second pair of curve generating processes allowing outliers (from now on, CGP2_{out}) with $s \in [0, 2\pi]$

$$x_0(s) = \begin{cases} u_{01} \sin s + u_{02} \cos s, & \text{with probability } 1 - q, \\ u_{01} \sin s + u_{12} \cos s, & \text{with probability } q. \end{cases}$$

$$x_1(s) = u_{11} \sin s + u_{12} \cos s,$$

where $u_{i,j}$, $i = 0, 1$, $j = 1, 2$ are defined as for CGP2 and $0 < q < 1$.

In Figure 3.2 we report a simulated dataset from CGP1, CGP2, CGP1_{out} and CGP2_{out} .

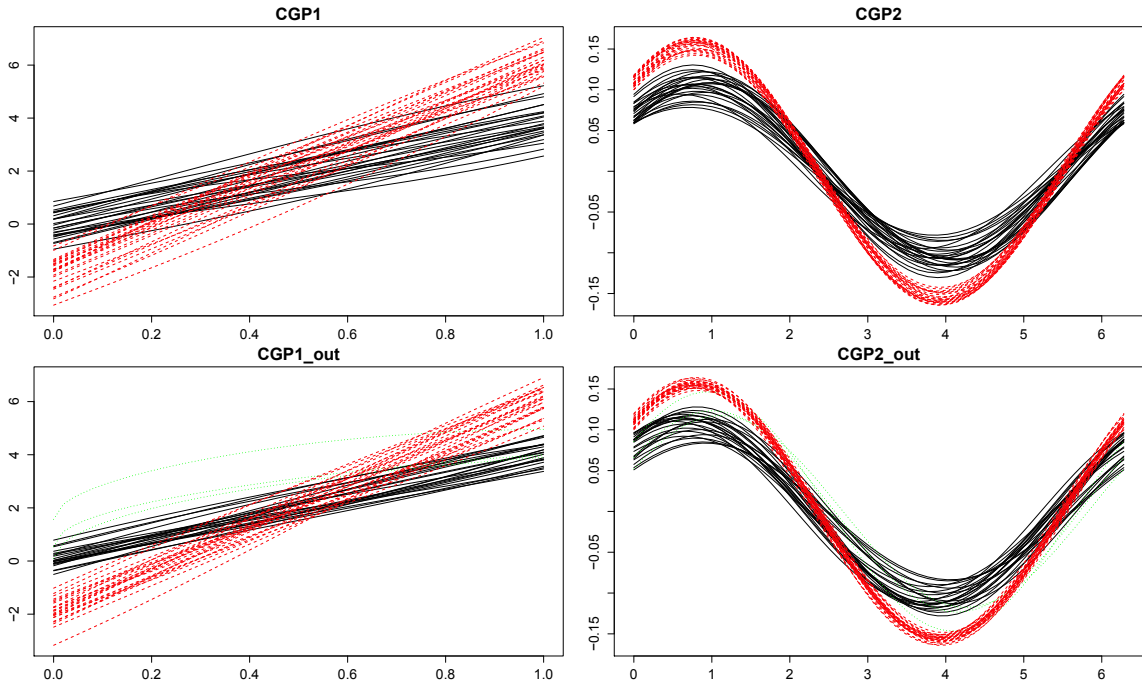


Figure 3.1: Simulated datasets from CGP1, CGP2, CGP1_{out} and CGP2_{out} : each dataset contains 25 curves from group $g = 0$ and 25 dashed curves from group $g = 1$. For CGP1 and CGP2, the curves from group $g = 0$ are all noncontaminated (solid). For CGP1_{out} and CGP2_{out} , the curves from group $g = 0$ can be noncontaminated (solid) or contaminated (dotted)

Next, we present some details of the simulation study: for each model, we generated 125 replications. For CGP1_{out} and CGP2_{out} , we set the contamination probability $q = 0.10$. For each replication, we generated 100 curves, 50 for $g = 0$ and 50 for $g = 1$. We used 25 curves from $g = 0$ and 25 curves from $g = 1$ to build each training sample, and we classified the

remaining curves. All curves were generated using a discretized and finite set of 51 equidistant points between 0 and 1 or 0 and 2π , depending on the model. For all the functional depths, we used a discretized version of their definitions. We performed the comparison among methods in terms of misclassification percentages. We report means and standard deviations of the misclassification percentages in Tables 3.1-3.4.

For what concerns KFSD, since we initially consider the set of percentiles $\{p_k\}$ to set σ , for each replication and classification method, we have based the choice of the appropriate percentile among the 7 options on a cross-validation step. Next, we describe it:

1. We divide each initial training sample in 5 groups-balanced cross-validation test samples, each one of size 10, and we pair each of them with its natural cross-validation training sample composed of the remaining 40 curves.
2. Therefore, for a each replication, we have available five pairs of cross-validation training and test samples. For each classification method and percentile, and for all the pairs, we classify the curves in the test samples and obtain a misclassification percentage.
3. For each replication and classification method, we search for the minimum misclassification percentage and its corresponding percentile, say p^* .

Then, for each replication and classification method, we use exclusively p^* to set the bandwidth of KFSD to do classification with the initial pair of training and test samples. In a preliminary stage of the study, since we observed several ties in the cross-validation step described above, we used the following second criteria to break the ties:

- for DTM, we select the percentile that minimizes the sum of the distances of the curves in the cross-validation test samples to the α -trimmed mean of the group at which the curves belong;
- for WAD, we select the percentile that minimizes the sum of the weighted averaged distances of the curves in the cross-validation test samples to the group at which the curves belong;

- for WMD, we select the percentile that maximizes the sum of the scaled KFSD of the curves in the cross-validation test samples when included in the group at which they belong.

The second criteria break almost all the ties. However, if after considering the second criterion a tie is not broken, we break it randomly.

Before showing the results, we discuss some computational issues. Note that for both DTM and WAD the key step consists in computing the within-group depth values of the training curves. Hence, we report the computational times (in seconds) that require each functional depth to perform this task for a training sample of size 50: 0.02 for FMD, 0.08 for HMD, 0.02 for RTD, 0.02 for IDD, less than 0.01 for MBD, 0.05 for FSD, and 0.14 for KFSD². Additionally, a 7-percentiles KFSD analysis takes 0.54 seconds, and a 7-percentiles KFSD analysis combined with a 5-subsamples cross-validation step takes 1.70 seconds. Therefore, the computation of KFSD is widely feasible, and the option of searching for an appropriate percentile by means of a cross-validation does not cause major computational problems, whereas it will bring some classification benefits.

Tables 3.1-3.4 report the performances observed with the data generated from CGP1, CGP2, CGP1_{out} and CGP2_{out}. Regarding the KFSD-based methods, for each pair of initial training and test samples and classification method, it may happen to observe the same misclassification error using the 7 different percentiles. In such cases, the cross-validation step turns out to be unnecessary. On the contrary, the cross-validation is required if there are at least two different misclassification errors. For this reason, for each model and method, in Table 3.5 we report the percentages of replications where the KFSD cross-validation step is required. In the same table we also report the best performing percentiles for the pairs of initial training and test samples since this information permits to identify the type of local analysis that would classify best each type of data.

²For FMD, HMD, RTD and IDD we have used the corresponding R functions that are available in the R package `fda.usc` on CRAN (Febrero and Oviedo de la Fuente 2012); for MBD we have followed the guidelines contained in Sun et al. (2012); for FSD and KFSD we have built some functions for R, which are available upon request. Features of the workstation: Intel Core i7-3.40GHz and 16GB of RAM.

Table 3.1: CGP1. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	0.16	0.22	0.14	0.14	0.18	0.19	0.14
	(0.65)	(0.73)	(0.58)	(0.63)	(0.62)	(0.64)	(0.58)
WAD	0.10	0.16	0.11	0.11	0.08	0.13	0.10
	(0.50)	(0.60)	(0.53)	(0.53)	(0.47)	(0.55)	(0.50)
WMD	15.09	1.66	21.90	18.06	11.82	3.30	0.13
	(5.43)	(2.46)	(6.82)	(6.27)	(4.95)	(2.89)	(0.66)
k -NN				0.11			
				(0.46)			

Table 3.2: CGP2. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	2.43	2.59	2.40	2.45	2.45	2.42	2.35
	(2.01)	(2.24)	(2.00)	(2.05)	(1.98)	(1.99)	(2.10)
WAD	3.38	3.33	3.10	3.22	3.20	3.02	3.14
	(2.32)	(2.35)	(2.25)	(2.33)	(2.23)	(2.19)	(2.25)
WMD	0.10	2.16	1.12	0.99	0.13	0.14	0.11
	(0.43)	(2.75)	(1.99)	(1.75)	(0.49)	(0.52)	(0.53)
k -NN				0.88			
				(1.53)			

Table 3.3: CGP1_{out}. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	0.11	0.08	0.08	0.08	0.06	0.11	0.05
	(0.46)	(0.39)	(0.39)	(0.39)	(0.35)	(0.46)	(0.31)
WAD	0.06	0.06	0.06	0.08	0.06	0.06	0.06
	(0.35)	(0.35)	(0.35)	(0.39)	(0.35)	(0.35)	(0.35)
WMD	14.46	2.34	22.93	18.93	11.47	3.26	0.08
	(5.54)	(2.97)	(6.92)	(7.11)	(5.02)	(2.80)	(0.39)
k -NN				0.08			
				(0.39)			

Table 3.4: CGP2_{out}. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	3.87	4.11	3.84	3.81	3.76	3.74	3.71
	(2.86)	(3.11)	(2.84)	(2.75)	(2.86)	(2.83)	(2.91)
WAD	4.94	4.94	4.67	4.74	4.80	4.48	4.70
	(3.36)	(3.44)	(3.25)	(3.25)	(3.24)	(3.16)	(3.29)
WMD	0.82	2.90	4.19	3.84	0.83	0.83	0.58
	(1.48)	(3.31)	(3.72)	(3.26)	(1.55)	(1.46)	(1.29)
k -NN				1.97			
				(2.05)			

Table 3.5: DTM+KFSD, WAD+KFSD and WMD+KFSD, and curve generating processes CGP1, CGP2, CGP1_{out} and CGP2_{out}: percentages of the replications for which cross-validation is required and best performing percentiles for the initial training and test samples

Model	Method	Percentage	Percentiles	Model	Method	Percentage	Percentiles
CGP1	DTM	7.20	50%	CGP2	DTM	21.60	66%
	WAD	2.40	15%, 25%, 50%, 66%, 75%		WAD	13.60	85%
	WMD	56.00	50%, 66%, 75%		WMD	12.80	66%, 85%
CGP1 _{out}	DTM	3.20	15%, 25%, 50%, 66%	CGP2 _{out}	DTM	26.40	85%
	WAD	0.00	all		WAD	19.20	85%
	WMD	58.40	66%		WMD	37.60	75%

The results in Tables 3.1-3.5 show that:

1. When the curves are generated from CGP1 or CGP1_{out}, WAD is the best classification procedure, but also the performances of DTM are competitive. Both procedures turn out rather stable with respect to the choice of the functional depth. On the contrary, WMD is more sensitive to the choice of the depth measure, and only the combination WMD+KFSD is able to compete with WAD and DTM. The performances of k -NN are quite good, but they are not the best ones neither for CGP1 nor for CGP1_{out}. Indeed, for CGP1, the best method is WAD+MBD (0.08%), whereas WAD+KFSD and WAD+FMD are the second best methods (0.10%). For CGP1_{out}, the best method is DTM+KFSD (0.05%), and some other spatial depth-based methods perform also quite good, e.g., WAD+FSD and WAD+KFSD (0.06%). Finally, note that WMD+KFSD behaves reasonably well in both scenarios (0.13% with CGP1 and 0.08% with CGP1_{out}).
2. When the curves are generated from CGP2 or CGP2_{out}, WMD, in conjunction with FMD, MBD, FSD and KFSD, is clearly the best performing classification procedure: indeed, the four resultant methods markedly outperform k -NN. For CGP2, the best method is WMD+FMD (0.10%), but the performance of WMD+KFSD is almost equal (0.11%). For CGP2_{out}, the best method is clearly WMD+KFSD (0.58%).
3. The cross-validation step is in general required more by WMD+KFSD, and to a smaller extent by DTM+KFSD, and finally by WAD+KFSD. For example, with CGP1_{out} the cross-validation step is completely unnecessary for WAD+KFSD, whereas it is advisable for

more than one-half of the replications with WMD+KFSD. On the other hand, looking at the best performing percentiles and focusing on the methods highlighted in the previous two points, we observe the following: under CGP1 and CGP1_{out}, WAD+KFSD reaches its best performances with several percentiles; under CGP2 and CGP2_{out}, the best performances of WMD+KFSD are with rather high percentiles. This last result is coherent with the good performances of WMD+FSD under CGP2 and CGP2_{out}. Indeed, the higher is the percentile, the less local-oriented is KFSD, and its behavior tends towards the behavior of FSD, its global-oriented counterpart.

Observing the curves generated from CGP1 in [Figure 3.1](#), it is possible to appreciate a rather strong data dependence structure due to the covariance function given by (3.1). On the other hand, observing the curves generated from CGP2 at any fixed $s \in [0, 2\pi]$, a low data variability stands out. We enhance our simulation study by relaxing these two features of CGP1 and CGP2 and considering two modifications of them. First, we consider a variation of CGP1 (from now on, CGP3) which consists in substituting the covariance function of the additive zero-mean Gaussian component previously given by (3.1) with a weaker version defined by

$$\mathbb{E}(\epsilon(s), \epsilon(s')) = 0.30 \exp(-|s - s'|/0.3), \quad s, s' \in [0, 1].$$

Second, we consider a variation of CGP2 (from now on, CGP4) which consists in adding to the two processes in (3.2) two identical additive zero-mean Gaussian components having covariance function

$$\mathbb{E}(\epsilon(s), \epsilon(s')) = 0.00025 \exp(-(s - s')^2), \quad s, s' \in [0, 2\pi].$$

[Figure 3.2](#) reports two simulated datasets from CGP3 and CGP4.

We use CGP3 and CGP4 to develop the third and last part of the simulation study. [Tables 3.6](#) and [3.7](#) report the performances of the 21 depth-based methods and of k -NN, whereas [Table 3.8](#) is the analogous of [Table 3.5](#) for CGP3 and CGP4.

The results in [Tables 3.6-3.8](#) show that:

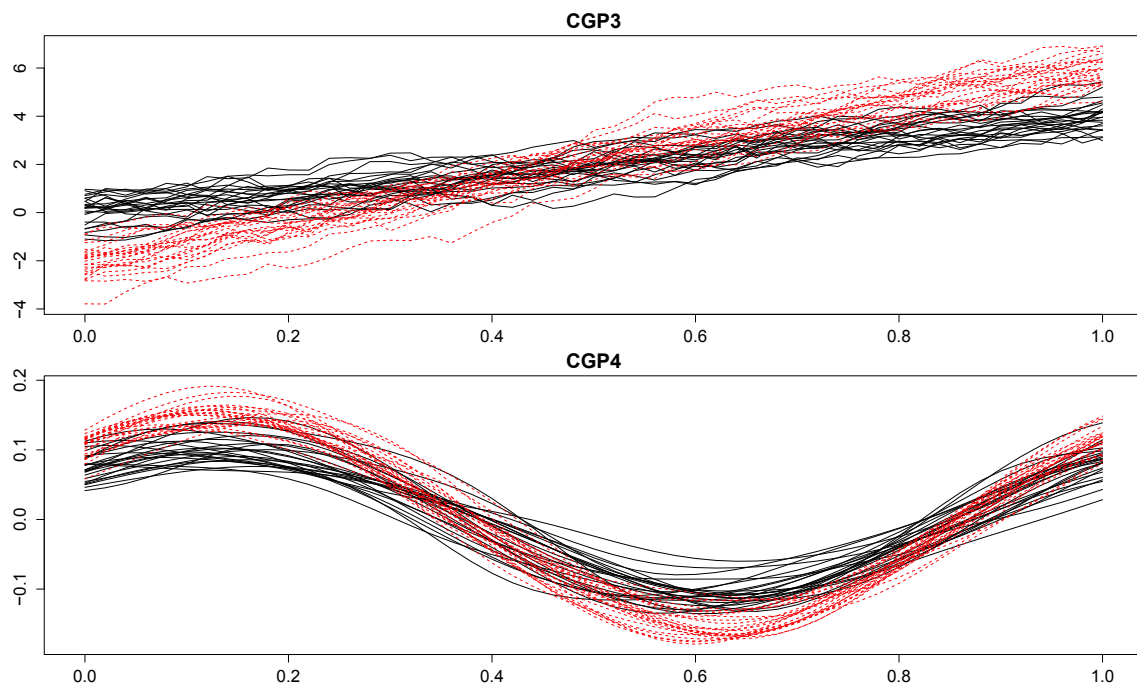


Figure 3.2: Simulated datasets from CGP3 and CGP4: each dataset contains 25 solid curves from group $g = 0$ and 25 dashed curves from group $g = 1$

Table 3.6: CGP3. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	1.33	1.36	1.39	1.36	1.36	1.38	1.31
	(1.39)	(1.40)	(1.44)	(1.47)	(1.31)	(1.42)	(1.35)
WAD	1.14	1.18	1.20	1.17	1.17	1.20	1.17
	(1.33)	(1.32)	(1.37)	(1.35)	(1.32)	(1.34)	(1.32)
WMD	5.26	2.93	17.42	14.64	4.29	1.44	1.20
	(3.12)	(2.66)	(5.90)	(5.51)	(2.76)	(1.56)	(1.39)
k -NN				1.39			
				(1.46)			

Table 3.7: CGP4. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	2.38	2.53	2.30	2.38	2.26	2.30	2.26
	(2.61)	(2.59)	(2.46)	(2.47)	(2.42)	(2.45)	(2.40)
WAD	3.55	3.46	3.46	3.46	3.60	3.36	3.41
	(3.08)	(3.08)	(2.87)	(2.98)	(3.07)	(2.92)	(2.94)
WMD	15.52	2.88	28.54	24.02	12.96	4.90	0.82
	(5.29)	(3.19)	(6.99)	(6.98)	(4.76)	(3.96)	(1.57)
k -NN				1.81			
				(2.28)			

Table 3.8: DTM+KFSD, WAD+KFSD and WMD+KFSD, and curve generating processes CGP3 and CGP4: percentages of the replications for which cross-validation is required and best performing percentiles for the initial training and test samples

Model	Method	Percentage	Percentiles	Model	Method	Percentage	Percentiles
CGP3	DTM	12.00	25%, 75%, 85%	CGP4	DTM	14.40	85%
	WAD	0.80	25%, 33%, 50%, 66%, 75%		WAD	7.20	66%
	WMD	28.00	75%, 85%		WMD	77.60	33%

1. When the curves are generated from CGP3, which is a modification of CGP1, we observe similar results. WAD is the best classification procedure, but the behavior of DTM is not bad at all. WMD heavily fails, with the exceptions of the spatial depths KFSD and FSD, and HMD. k -NN is a competitive classification procedure, but all the DTM and WAD-based methods, and WMD+KFSD outperform it. The best method is WAD+FMD (1.14%), whereas there are several best second methods (1.17%), including WAD+KFSD.
2. When the curves are generated from CGP4, which is a modification of CGP2, we observe that WMD+KFSD is the only method able to outperform k -NN (0.82% against 1.81%), whereas the remaining methods highlighted for CGP2, i.e., WMD+FMD, WMD+MBD and WMD+FSD, drastically worsen.
3. As for the previous models, the cross-validation step is in general required more by WMD+KFSD, and to a smaller extent by DTM+KFSD, and finally by WAD+KFSD. For example, with CGP3 the cross-validation step is almost unnecessary for WAD+KFSD, whereas it is advisable for more than one-quarter of the replications with WMD+KFSD. On the other hand, looking at the best performing percentiles and focusing on the methods highlighted in the previous two points, we observe that for CGP3 WAD+KFSD reaches its best performance with all the percentiles except the 15% one, whereas for CGP4 the best performances of WMD+KFSD are with the 33% percentile, which means that for this method there is a gain when a rather strong local approach is implemented.

To conclude, we have observed that for the curve generating processes having a deterministic and linear mean function and a random component, i.e., CGP1, CGP1_{out} and CGP3, WAD+KFSD is among the best and most stable classification methods, and both DTM+KFSD

and WMD+KFSD have performances that are not so different. On the other hand, for the curve generating processes having a random and nonlinear mean function, i.e., CGP2, CGP2_{out} and CGP4, WMD+KFSD is clearly the best classification method. Therefore, KFSD-based functional supervised classification is certainly a good option to discriminate curves.

3.4 Real Data Study

To complete the comparison among the depth-based supervised classification methods and k -NN, we also consider two real datasets.

3.4.1 Growth Data

The first real dataset consists of 93 growth curves: 54 are heights of girls, 39 are heights of boys. All of them are observed at a common discretized set of 31 nonequidistant ages between 1 and 18 years. [Figure 3.3](#) shows the curves (for more details about this dataset, see [Ramsay and Silverman 2005](#)). We use natural cubic spline interpolation to estimate the growth curves at a common and equally spaced domain.

This dataset has been analyzed by [López-Pintado and Romo \(2006\)](#) and [Cuevas et al. \(2007\)](#). From our point of view, these data are interesting mainly for two reasons: first, the differences between the two groups are not so much sharp; and second, we can not discard the presence of some outlying curves, especially among girls.

We perform the first part of the growth data classification study with a similar structure to the simulation study. More precisely, we consider 150 training samples composed of 40 and 30 randomly chosen curves of girls and boys, respectively. We pair each training sample with the test sample composed of the remaining 14 and 9 curves of girls and boys, respectively. We denote this way of obtaining training and test samples as T1, and we try to classify the curves included in each test sample by using the methods and depths considered in [Section 3.4](#), with the same specifications for both.

For what concerns the cross-validation step for KFSD, we divide each initial training sam-

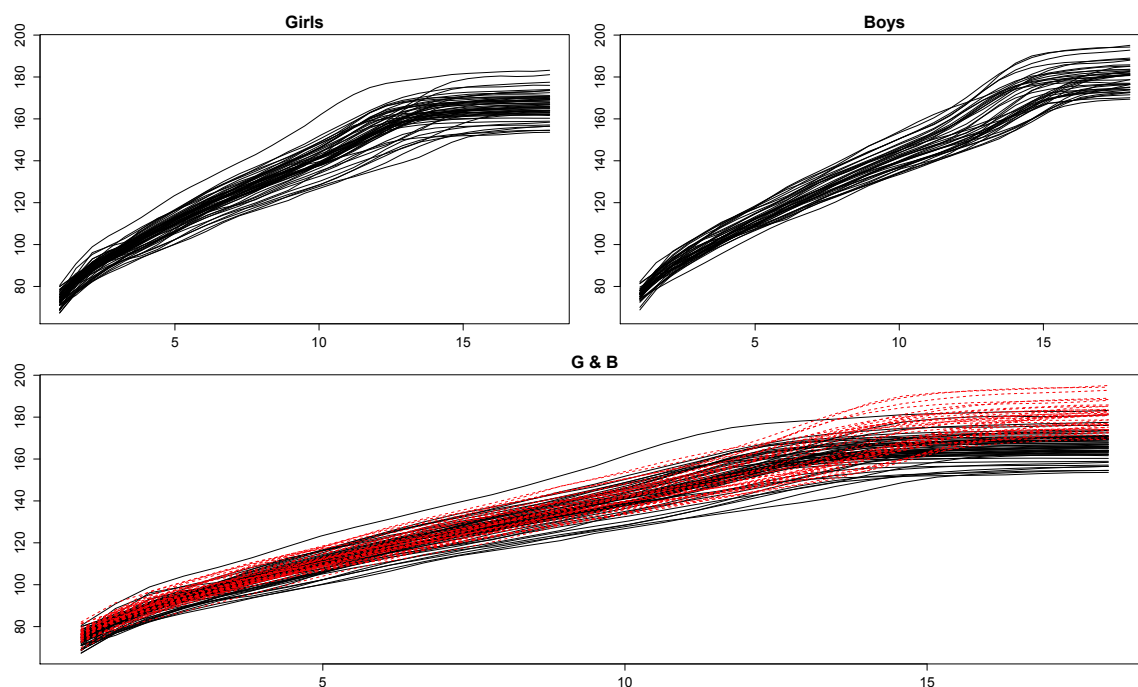


Figure 3.3: Growth curves: 54 heights of girls (top left), 39 heights of boys (top right), 93 heights of girls and boys (bottom; the dashed curves are boys)

ple in 5 groups-unbalanced cross-validation test samples with 8 and 6 curves of girls and boys, respectively, and we pair each of them with its natural cross-validation training sample.

We report the performances of the 21 depth-based methods and k -NN in Table 3.9. In Table 3.11 we report the percentages of pairs of initial training and test samples for which the cross-validation step is required by DTM+KFSD, WAD+KFSD and WMD+KFSD, and we also show the best performing percentiles for the same samples.

Table 3.9: Growth data and T1. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	15.22 (8.18)	10.84 (7.35)	19.33 (8.78)	19.91 (8.99)	17.30 (8.80)	19.45 (9.08)	12.14 (7.82)
WAD	14.43 (7.56)	11.22 (7.41)	15.10 (8.18)	15.16 (7.88)	14.87 (8.11)	15.42 (8.12)	12.84 (7.48)
WMD	30.41 (11.31)	4.96 (4.56)	35.36 (8.60)	33.16 (8.79)	27.59 (10.22)	18.03 (7.39)	3.45 (3.57)
k -NN				3.86 (3.56)			

Additionally, we consider important to study how the classification methods work when

the goal is to classify a single curve using the information contained in the rest of the curves. To do this with the growth data, we consider each possible training sample composed of 92 curves, and classify the curve not included in the training set. We denote this way of obtaining training and test samples as T2, and we implement a cross-validation step for DTM+KFSD, WAD+KFSD and WMD+KFSD also for T2. We report the performances of the classification methods under T2 in Tables 3.10 and 3.11.

Table 3.10: Growth data and T2. Number of misclassified curves with DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	15	9	17	19	15	18	11
WAD	12	10	13	13	13	13	11
WMD	28	3	32	30	24	16	2
k -NN	3						

Table 3.11: Growth data and DTM+KFSD, WAD+KFSD and WMD+ KFSD. Percentages of initial T1-type and T2-type training and test samples for which cross-validation is required and best performing percentiles for the same samples

T	Method	Percentage	Percentiles	T	Method	Percentage	Percentiles
T1	DTM	72.00	33%	T2	DTM	3.23	25%
	WAD	36.67	50%		WAD	1.08	50%, 66%
	WMD	91.33	15%		WMD	10.75	15%

The results in Tables 3.9-3.11 show that WMD+KFSD is the only method able to outperform k -NN, an occurrence that has been already observed with curves generated from CGP4. Indeed, under T1, WMD+KFSD outperforms k -NN in terms of means of the misclassification percentages (3.45% against 3.86%), whereas the third best method is WMD+HMD (4.96%); something similar happens under T2: WMD+KFSD misclassifies 2 curves, and it is still the best method, followed by k -NN and WMD+HMD, which misclassify 3 curves. If we convert these T2 performances in percentages, i.e., $\left(\frac{\#\text{misclassified curves}}{\text{sample size}} \times 100\right)$, we observe that, moving from T1 to T2, there is a slight but systematic improvement: WMD+KFSD, 3.45% \rightarrow 2.16%; k -NN, 3.86% \rightarrow 3.23%; WMD+HMD, 4.96% \rightarrow 3.23%. This pattern is due to the greater size of the training samples under T2, but it does not cause significant changes in the performances-based order of the methods.

For what concerns the cross-validation step of the KFSD-based classification methods, most

of the remarks made for the simulated data also hold for this dataset. Moreover, it is clear that the implementation of the cross-validation step is a key issue under T1, whereas it becomes much less important under T2. Finally, under both T1 and T2, the best performing percentile for the best method, i.e., WMD+KFSD, is the 15% percentile, which means that classification of growth curves requires a strongly local approach.

Even though here we do not show the results obtained with the 7 different percentiles, we would like to report that, when we combine WMD with KFSD and use a fixed percentile, higher percentiles make the performances of the classification method worse. For example, using WMD and KFSD with the 15% and the 25% percentile, under T1 we observe means equal to 3.68% and 4.70%, respectively, whereas under T2 the methods misclassify 3 and 4 curves, respectively. Given the results under T2, we look at the misclassified curves by these two versions of WMD+KFSD and k -NN: using the 15% percentile, WMD+KFSD misclassifies girls with labels 11, 25 and 49; using the 25% percentile, it misclassifies girls with labels 8, 25, 49 and 38; k -NN misclassifies girls with labels 8, 25 and 49 (see [Figure 3.4](#)).

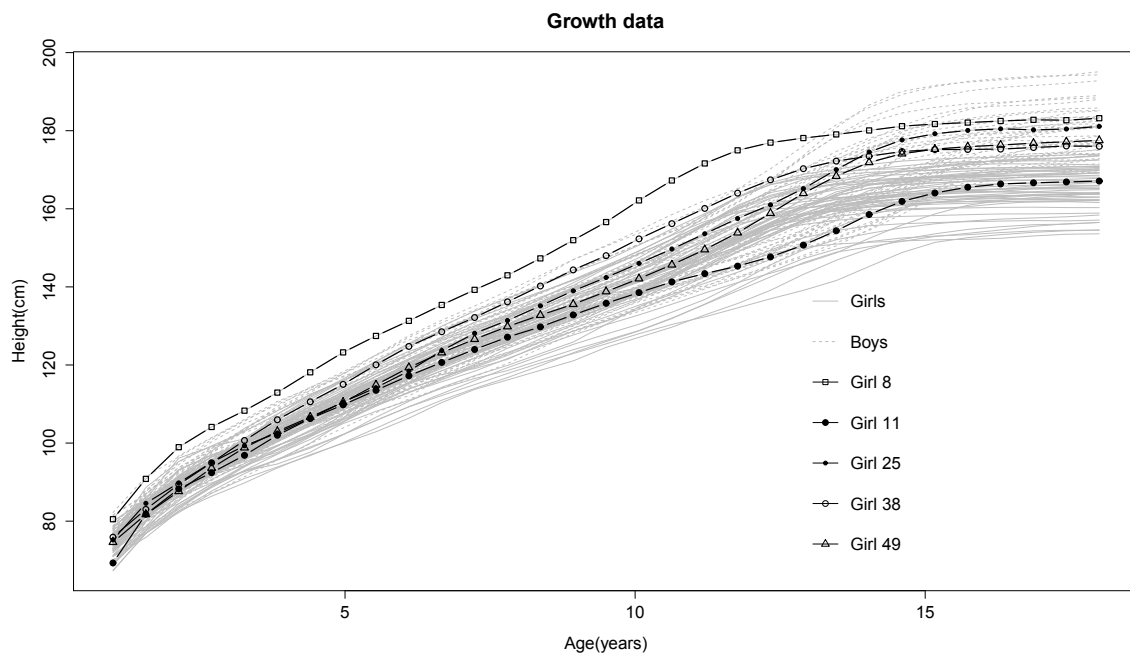


Figure 3.4: Growth curves: highlighting some interesting curves for the classification problem

Therefore, the differences between WMD+KFSD when used with the 15% percentile and k -NN lie in girls 8 and 11. Observing these curves, we can appreciate that with a local spatial approach it is possible to classify correctly a female height having apparently an outlying behavior (Girl 8), however at the price of misclassifying a more central female height (Girl 11); on the contrary, k -NN makes the opposite, and its behavior is more similar to the behavior of WMD+KFSD when used with the 25% percentile, which misclassifies the same curves as k -NN, in addition to the girl with label 38. Thanks to the cross-validation step, which allows a non-fixed percentile, WMD+KFSD takes advantage of the differences between the use of the 15% and the 25% percentile, and it succeeds in misclassifying only girls with labels 25 and 49.

3.4.2 Phoneme Data

The second real dataset that we consider consists in log-periodograms of length 150 corresponding to recordings of speakers pronouncing the phonemes “aa” or “ao”. More precisely, the dataset contains 400 recordings of the phoneme “aa” and 400 recordings of the phoneme “ao”. Since we are considering a large number of methods, we perform the study using 100 randomly chosen recordings of the phoneme “aa” (from now on, AA curves) and 100 randomly chosen recordings of the phoneme “ao” (from now on, AO curves). [Figure 3.5](#) shows the curves. For more details about this dataset, see [Ferraty and Vieu \(2006\)](#).

Observing [Figure 3.5](#), we can appreciate similar features to the ones highlighted for the growth curves: first, we can not discard the presence of some outlying curves in both groups; and second, the differences between the two groups are not so much sharp. Indeed, this second feature seems exaggerated in the second part of the data (frequencies from 76 to 150), and the discriminant information seems to lie especially in the first part of them (frequencies from 1 to 75). This hypothesis has been confirmed by a preliminary classification analysis in which we have observed that in general any method improves its performances when only the first half of the curves is used. Then, we perform the phoneme data classification study using the first 75 frequencies, and with a structure that is similar to the one used for the growth data, that is, we classify curves included in test samples that we obtain by means of both T1 and T2.

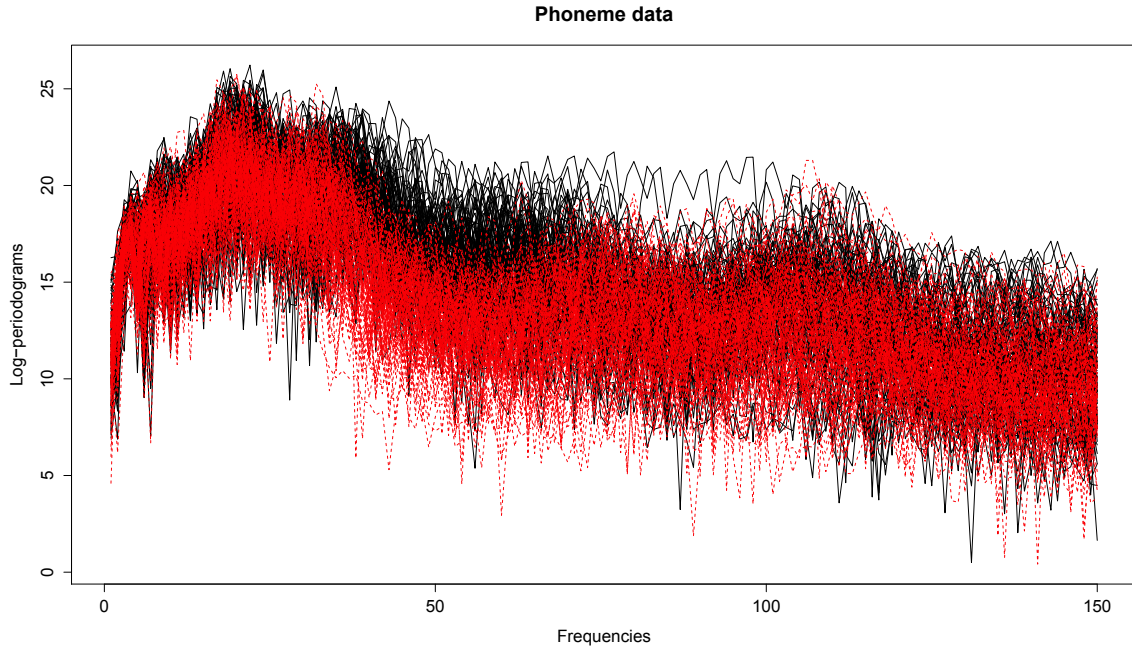


Figure 3.5: Phoneme data: log-periodograms of 100 AA curves (solid) and 100 AO curves (dashed)

To perform the T1 part of the study, we consider 100 training samples composed of 75 randomly chosen AA curves and 75 randomly chosen AO curves. Each training sample is paired with the test sample composed of the remaining 25 AA curves and 25 AO curves (i.e., the allocation “150 training curves, 50 test curves” defines T1), and we classify the curves included in each test sample by using the same methods and depths as in [subsection 3.4.1](#), with the same specifications for both.

For what concerns the cross-validation step for KFSD, it is similar to the one implemented for the simulated data, but in this case we divide each initial training sample in 5 groups-balanced cross-validation test samples of size 30 and we pair each of them with its natural cross-validation training sample of size 120.

We report the performances of the 21 depth-based methods and k -NN in [Table 3.12](#), whereas [Table 3.14](#) is the analogous of [Table 3.11](#) for the phoneme data.

To perform the T2 part of the study, we consider all the possible 200 training samples composed of 199 phonemes, each one jointly with its corresponding test sample composed

Table 3.12: Phoneme data and T1. Means and standard deviations (*in parenthesis*) of the misclassification percentages for DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	21.56 (5.07)	23.16 (5.64)	22.68 (5.47)	22.70 (5.37)	21.84 (5.18)	23.00 (5.56)	23.08 (5.60)
WAD	23.12 (5.49)	23.74 (5.82)	23.88 (5.78)	23.84 (5.73)	23.54 (5.60)	23.64 (5.82)	23.36 (5.78)
WMD	21.42 (4.56)	24.76 (5.50)	26.18 (5.79)	25.62 (6.06)	20.54 (4.57)	20.62 (4.86)	19.30 (4.66)
k -NN				22.14 (5.01)			

of the remaining curve. As for the growth data, we implement a cross-validation step for DTM+KFSD, WAD+KFSD and WMD+KFSD also under T2. We report the performances of the classification methods under T2 in Tables 3.13 and 3.14.

Table 3.13: Phoneme data and T2. Number of misclassified curves with DTM, WAD and WMD-based methods and k -NN

Method/Depth	FMD	HMD	RTD	IDD	MBD	FSD	KFSD
DTM	43	46	45	46	44	45	46
WAD	46	50	51	52	46	49	46
WMD	43	51	50	51	39	39	37
k -NN				45			

Table 3.14: Phoneme data and DTM+KFSD, WAD+KFSD and WMD+KFSD. Percentages of initial T1-type and T2-type training and test samples for which cross-validation is required and best performing percentiles for the same samples

T	Method	Percentage	Percentiles	T	Method	Percentage	Percentiles
T1	DTM	10.00	25%	T2	DTM	0.00	15%, 25%, 33%, 50%, 66%, 75%, 85%
	WAD	33.00	15%		WAD	1.00	15%, 25%, 33%, 50%, 66%
	WMD	54.00	15%		WMD	1.50	15%, 25%, 33%, 50%, 66%, 75%

The results in Tables 3.12-3.14 show especially two facts: first, classification of phoneme data is a hard problem, and effectively the number of misclassified curves is large with any method and under both T1 and T2; second, WMD+KFSD is the best classification method. Indeed, under T1, WMD+KFSD is the method with the best performance in terms of mean of the misclassification percentages (19.30%), and it outperforms the second best method, WMD+MBD (20.54%). The third best method is given by another spatial depth-based method, WMD+FSD (20.62%). Under T2, WMD+KFSD is again the method with the best performance

(37 misclassified curves), whereas WMD+MBD and WMD+FSD are the second best methods (39 misclassified curves). Note that the performance of the fourth best method is quite distant (WMD+FMD, 43 misclassified curves), as well as the one of k -NN (45 misclassified curves). If we convert the T2 performances of the three best methods in percentages, there is a slight but systematic improvement when moving from T1 to T2: WMD+KFSD, 19.30% \rightarrow 18.50%; WMD+MBD, 20.54% \rightarrow 19.50%; WMD+FSD, 20.62% \rightarrow 19.50%. However, as for growth data, we observe no significant changes in the performances-based order of the methods.

Observing [Table 3.14](#) and focusing on the best KFSD-based method, i.e., WMD+KFSD, we appreciate that under T1 the best performing percentile for WMD+KFSD is the 15% percentile, whereas for T2 the best performing percentiles for WMD+KFSD are all except the 85% percentile. However, even though we do not show the results obtained with the 7 different percentiles, we would like to report that for the phoneme data, unlike for the growth data, when we combine WMD with KFSD and a fixed percentile, even with the worst performing percentile, which is the 85% percentile, WMD+KFSD has performances comparable to the best ones. Indeed, using the 85% percentile, WMD+KFSD still outperforms the second best method of [Table 3.12](#) (19.90% against 20.54% of WMD+MBD under T1), and it misclassifies the same number of curves as the second best methods of [Table 3.13](#) (39 misclassified curves by WMD+MBD and WMD+FSD under T2).

3.5 Conclusions

After presenting KFSD in [Chapter 2](#), in this chapter we focused on supervised functional classification problems using the depth-based methods DTM, WAD and WMD, and a benchmark procedure such as k -NN. The three depth-based methods were used together with KFSD, FSD and five more existing functional depths. We developed a simulation study where the differences between curves of different groups are not excessively marked and/or the data may contain outliers, and we analyzed two real datasets with similar features. In both cases, we observed that a KFSD-based method is always among the best methods in terms of classifica-

tion capabilities and that it outperforms the benchmark procedure k -NN. Note that no other depth behaved as well as KFSD, and that WMD+KFSD produced doubtless the most stable and best depth-based classification method: indeed, WMD+KFSD has always outperformed WMD+FSD, which is its natural global-oriented competitor, and, more in general, it had acceptable results and often the best ones, as in the case of growth and phoneme data.

*Io so. Ma non ho le prove.
Non ho nemmeno indizi.
(Pier Paolo Pasolini, Cos' è questo golpe? Io so)*

Chapter 4

Functional Outlier Detection¹

4.1 Introduction

The accurate identification of outliers is an important aspect in any statistical data analysis. Nowadays there are well-established outlier detection techniques in the univariate and multivariate frameworks (for a complete review, see for example [Barnett and Lewis 1994](#)) as well as in FDA. The possible reasons why statisticians are interested in identifying outliers are mainly two: first, any statistical technique may lead to misleading results and conclusions when applied to contaminated datasets; second, since outliers may arise because of extraordinary deviations from the average pattern of the data, a further non-statistical analysis of extreme realizations can reveal interesting but latent features of the phenomenon under study.

According to [Febrero et al. \(2007, 2008\)](#), a functional outlier is a curve generated by a stochastic process with a different distribution than the distribution of the normal curves. This definition covers many types of outliers, e.g., magnitude outliers, shape outliers and partial outliers, i.e., curves having atypical behaviors only in some segments of the domain. Shape and partial outliers are typically harder to detect than magnitude outliers (in the case of high magnitude, outliers can even be recognized by simply looking at a graph), and therefore entail more difficult outlier detection problems. In this chapter we focus on such scenarios and, from

¹This chapter is mostly based on the working paper by [Sguera et al. \(2014a\)](#)

now on, we refer to low magnitude, shape and partial outliers as “faint outliers” and to high magnitude outliers as “clear outliers”.

When the local spatial depth KFSD is used to order sample curves from the most to the least central, in general it succeeds in ranking faint outliers among the least deep curves. To exploit at the best this feature of KFSD, we introduce three new procedures that provide a threshold value for KFSD such that curves with depth values lower than the threshold are detected as outliers: first, we present a result that allows to select a threshold for KFSD to detect outliers. This result is based on a probabilistic upper bound on a desired false alarm probability of detecting normal curves as outliers. However, its practical application requires the availability of two samples, circumstance rather uncommon in classical outlier detection problems. Then, we propose three solutions based on smoothed resampling techniques that require instead a unique sample.

We study the performances of the resampling-based procedures in a simulation study and in a real application with environmental data, and we compare them with the performances of some competitors. First, we consider the methods proposed by [Febrero et al. \(2008\)](#), who also suggested to label as outliers those curves with depth values lower than a certain threshold: as functional depths, they considered three alternatives, i.e., FMD, HMD and IDD, whereas, they proposed two alternative bootstrap procedures based on depth-based trimmed and weighted resampling, respectively, to determine the depth threshold. Second, we employ the functional boxplot presented by [Sun and Genton \(2011\)](#). The proposed functional boxplot is constructed using the ranking of curves provided by MBD and allows to detect outliers as well as the standard boxplot does. Finally, since the use of a functional depth is only one among the possible strategies for tackling the functional outlier detection problem, we also consider the methods proposed by [Hyndman and Shang \(2010\)](#), who first reduce the outlier detection problem from functional to multivariate data by means of functional principal component analysis (FPCA), and then use two alternative multivariate techniques on the scores to detect outliers, i.e., the bagplot and the high density region boxplot, respectively. The results of the comparative study support our proposals.

4.2 Outlier Detection for Functional Data

The outlier detection problem can be described as follows: let $Y_n = \{y_1, \dots, y_n\}$ be a sample that has been generated from a mixture of two functional random variables in \mathbb{H} , one for normal curves and one for outliers, say Y_{nor} and Y_{out} , respectively. Let Y_{mix} be this mixture, i.e.,

$$Y_{mix} = \begin{cases} Y_{nor}, & \text{with probability } 1 - \alpha, \\ Y_{out}, & \text{with probability } \alpha, \end{cases} \quad (4.1)$$

where $\alpha \in [0, 1]$ is the contamination probability (usually, a value rather close to 0). The curves composing Y_n are all unlabeled, and the goal of the analysis is to decide whether each curve is a normal curve or an outlier.

Since any functional depth measures the degree of centrality/extremality of a given curve relative to a distribution or a sample, outliers are expected to have low depth values. In [Chapter 3](#) we used depth-based methods to solve supervised functional classification problems. It was observed that a local approach is preferable when the classes involved in the problem are not extremely different or distant. Here, we show that an approach based on a local depth such as KFSD also succeeds in detecting faint outliers such as low magnitude, shape or partial outliers.

Recall that KFSD is a functional extension of the kernelized spatial depth for multivariate data (KSD) proposed by [Chen et al. \(2009\)](#), who also proposed a KSD-based outlier detector that we generalize to KFSD: for a given dataset Y_n generated from Y_{mix} and $t, b \in [0, 1]$, the KFSD-based outlier detector for $x \in \mathbb{H}$ is given by

$$g(x, Y_n) = \begin{cases} 1, & \text{if } KFSD(x, Y_n) \leq t, \\ \frac{t+b-KFSD(x, Y_n)}{b}, & \text{if } t < KFSD(x, Y_n) \leq t + b, \\ 0, & \text{if } KFSD(x, Y_n) > t + b, \end{cases} \quad (4.2)$$

where t is a threshold and b determines the transition rate of x from being an outlier (case $g(x, Y_n) = 1$) to be a normal curve (case $g(x, Y_n) = 0$). Clearly, (4.2) depends on the values of

t and b . On the one hand, it is desirable a value of t capable of discriminating between x generated from Y_{nor} or Y_{out} . On the other hand, the role of b depends on the goal of the analysis. If the options “outlier” and “normal curve” are the only ones of interest, b should be set at 0. However, if there is interest in further analysis of “potential outliers”, b may be allowed to be greater than 0. In our case, since the main goal is outlier detection and t is the key parameter to be set, we let $b = 0$.

For the multivariate case, [Chen et al. \(2009\)](#) studied KSD-based outlier detection under different scenarios. One of them consists in an outlier detection problem where two samples are available, and for which they proposed to select the threshold t by controlling the probability that normal observations are classified as outliers, i.e., the false alarm probability (FAP). They proved a result providing a KSD-based probabilistic upper bound on the FAP which depends on t . Then, the maximum value of t such that the upper bound does not exceed a given desired FAP provides a threshold for KSD. Next, we extend this result to KFSD:

Theorem 4.1. *Let $Y_{n_Y} = \{y_i, \dots, y_{n_Y}\}$ and $Z_{n_Z} = \{z_i, \dots, z_{n_Z}\}$ be two i. i. d. samples generated from the unknown mixture of random variables $Y_{mix} \in \mathbb{H}$ described by (4.1), with $\alpha > 0$. Let $g(\cdot, Y_{n_Y})$ be the outlier detector defined in (4.2). Fix $\delta \in (0, 1)$ and suppose that $\alpha \leq r$ for some $r \in [0, 1]$. For a new random element x generated from Y_{nor} , the following inequality holds with probability at least $1 - \delta$:*

$$\mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})] \leq \frac{1}{1-r} \left[\frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y}) + \sqrt{\frac{\ln 1/\delta}{2n_Z}} \right], \quad (4.3)$$

where $\mathbb{E}_{x \sim Y_{nor}}$ refers to the expected value with respect to x generated from Y_{nor} .

The proof of [Theorem 4.1](#) is presented in the appendix of this chapter. Recall that the FAP has been defined as the probability that a normal observation x is classified as outlier. For the elements of [Theorem 4.1](#), $\Pr_{x \sim Y_{nor}} (g(x, Y_{n_Y}) = 1)$ is the FAP. If we set $b = 0$,

$$\Pr_{x \sim Y_{nor}} (g(x, Y_{n_Y}) = 1) = \mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})].$$

Therefore, the probabilistic upper bound of [Theorem 4.1](#) applies also to the FAP.

It is worth noting that the application of [Theorem 4.1](#) requires to observe two samples, circumstance rather uncommon in classical outlier detection problems, but also a considerably large n_Z . To show the last point, recall that the right-hand side of (4.3) has to be controlled under the desired FAP and is in practice composed by two addends, with the second equal to $\frac{1}{1-r} \sqrt{\frac{\ln 1/\delta}{2n_Z}}$. For some normal values such as $r = 0.05$, $\delta = 0.05$ and $n_Z = 50$, we would have $\frac{1}{1-r} \sqrt{\frac{\ln 1/\delta}{2n_Z}} = 0.18$, which is greater than a normal desired FAP such as 0.10, and it shows that the use of [Theorem 4.1](#) may be compromised under some common situations.

We propose three solutions to overcome these limitations. Assume to observe a functional sample Y_n generated from an unknown mixture of random variables Y_{mix} . The goal is to identify which curves in Y_n are outliers, but in this situation there are not available two samples and [Theorem 4.1](#) cannot be applied. We propose to use Y_n as Y_{n_Y} , and to obtain Z_{n_Z} by resampling with replacement from Y_n . In this way, we also solve the problematic issue related to the second addend of the right-hand side of (4.3) because it is possible to set n_Z as large as needed. Regarding the resampling procedure to obtain Z_{n_Z} , we consider three different schemes, all of them with replacement. Since we deal with potentially contaminated datasets, besides simple resampling, we also consider two robust KFSD-based resampling procedures inspired by the work of [Febrero et al. \(2008\)](#). Then, the three resampling schemes that we consider are:

1. Simple resampling.
2. KFSD-based trimmed resampling: once obtained $KFSD(y_i, Y_n), i = 1, \dots, n$, it is possible to identify the $\lceil \alpha_T n \rceil\%$ of least deep curves, for certain $0 < \alpha_T < 1$ usually close to 0. These least deep curves are deleted from the sample, and simple resampling is carried out with the remaining curves.
3. KFSD-based weighted resampling: once obtained $KFSD(y_i, Y_n), i = 1, \dots, n$, weighted resampling is carried out with weights $w_i = KFSD(y_i, Y_n)$.

All the above procedures generate samples with some repeated curves. However, in a preliminary stage of our study we observed that it is preferable to work with Z_{n_Z} composed of curves different among them. To obtain such samples, we add a common smoothing step to

the previous three resampling schemes.

To describe the smoothing step, first recall that each curve in Y_n is in practice observed at a discretized and finite set of domain points, and that the sets may differ from one curve to another. For this reason, the estimation of Y_n at a common set of m equidistant domain points may be required. Let $(y_i(s_1), \dots, y_i(s_m))$ be the observed or eventually estimated m -dimensional equidistant discretized version of y_i , Σ_{Y_n} be the covariance matrix of the discretized form of Y_n and γ be a smoothing parameter. Consider a zero-mean Gaussian process whose discretized form has $\gamma\Sigma_{Y_n}$ as covariance matrix. Let $(\zeta(s_1), \dots, \zeta(s_m))$ be a discretized realization of the previous Gaussian process. Consider any of the previous three resampling procedures and assume that at the j th trial, $j = 1, \dots, n_Z$, the i th curve in Y_n has been sampled. Then, the discretized form of the j th curve in Z_{n_Z} would be given by $(z_j(s_1), \dots, z_j(s_m)) = (y_i(s_1) + \zeta_j(s_1), \dots, y_i(s_m) + \zeta_j(s_m))$, or, in functional form, by $z_j = y_i + \zeta_j$. Therefore, combining each resampling scheme with this smoothing step, we provide three different approximate ways to obtain Z_{n_Z} , and we refer to them as *smo*, *tri* and *wei*, respectively. Then, for fixed δ , r and desired FAP, the threshold t for (4.2) is selected as the maximum value of t such that the right-hand side of (4.3) does not exceed the desired FAP. Let t^* be the selected threshold, which is then used in (4.2) with $b = 0$ to compute $g(y_i, Y_n)$, $i = 1, \dots, n$. If $g(y_i, Y_n) = 1$, y_i is detected as outlier. To summarize, we provide three KFSD-based outlier detection procedures and we refer to them as KFSD_{smo} , KFSD_{tri} and KFSD_{wei} depending on how Z_{n_Z} is obtained (*smo*, *tri* and *wei*, respectively; recall that $Y_{n_Y} = Y_n$). As competitors of the proposed procedures, we consider the methods mentioned in Section 4.1 that we next describe.

[Sun and Genton \(2011\)](#) proposed a depth-based functional boxplot and an associated outlier detection rule based on the center-outward ordering of the sample curves provided by MBD. Once obtained the ordering, it is used to define the sample central region, that is, the smallest band containing at least half of the deepest curves. The non-outlying region is defined inflating the central region by 1.5 times. Curves that do not belong completely to the non-outlying region are detected as outliers. The original functional boxplot is based on the use of MBD, but clearly any functional depth can be used. Another contribution of this the-

sis is the study of the performances of the outlier detection rule associated to the functional boxplot (from now on, FBP) when used together with the battery of functional depths that we considered in supervised classification.

Febrero et al. (2008) proposed two depth-based outlier detection procedures selecting a threshold for FMD, HMD or IDD by means of two alternative robust smoothed bootstrap procedures whose single bootstrap samples are obtained using the above described *tri* and *wei*, respectively. Then, at each bootstrap sample, the 1% percentile of the empirical distribution of the depth values is obtained, say $p_{0.01}$. If B is the number of bootstrap samples, B values of $p_{0.01}$ are obtained. Each method selects as cutoff c the median of the collection of $p_{0.01}$ and, using c as threshold, a first outlier detection is performed. If some curves are detected as outliers, they are deleted from the sample, and the procedure is repeated until no more outliers are found (note that c is computed only in the first iteration). From now on, we refer to these methods as B_{tri} and B_{wei} . Also in this case, we evaluate these procedures using all the functional depths considered in Chapter 3.

Finally, we also consider two procedures that are not based on the use of a functional depth proposed by Hyndman and Shang (2010). Both procedures are based on the first two robust functional principal components scores and on two different graphical representations of them. The first proposal is the outlier detection rule associated to the functional bagplot (from now on, FBG), which works as follows: obtain the bivariate robust scores and order them using the multivariate halfspace depth HSD. Define an inner region by considering the smallest region containing at least the 50% of the deepest scores, and obtain a non-outlying region by inflating the inner region by 2.58 times. FBG detects as outliers those curves whose scores are outside the non-outlying region. Note that the scores-based regions and outliers allow to draw a bivariate bagplot, which produces a functional bagplot once it is mapped onto the original functional space. The second proposal is related to a different graphical tool, the high density region boxplot (from now on, we refer to its associated outlier detection rule as FHD). In this case, once obtained the scores, perform a bivariate kernel density estimation. Define the $(1 - \beta)$ -high density region (HDR), $\beta \in (0, 1)$, as the region of scores with coverage probability

equal to $(1 - \beta)$. FHD detects as outliers those curves whose scores are outside the $(1 - \beta)$ -HDR. In this case, it is possible to draw a bivariate HDR boxplot which can be mapped onto a functional version, thus providing the functional HDR boxplot.

4.3 Simulation Study

After introducing KFSD_{smo} , KFSD_{tri} and KFSD_{wei} and their competitors (FBP, B_{tri} , B_{wei} , FBG and FHD), in this section we carry out a simulation study to evaluate the performances of the different methods. Since we employ FBP, B_{tri} and B_{wei} with seven different functional depths (FMD, HMD, RTD, IDD, MBD, FSD and KFSD), for them we use the notation procedure+depth: for example, FBP+FMD refers to the method obtained by using FBP together with FMD.

We consider six models for our simulation study: all of them generate curves according to the mixture of random variables Y_{mix} described in (4.1). The first three mixture models (MM1, MM2 and MM3) share Y_{nor} , with curves generated by

$$y(s) = 4s + \epsilon(s), \quad (4.4)$$

where $s \in [0, 1]$ and $\epsilon(s)$ is a zero-mean Gaussian component with covariance function given by

$$\mathbb{E}(\epsilon(s), \epsilon(s')) = 0.25 \exp(-(s - s')^2), \quad s, s' \in [0, 1].$$

Also the remaining three mixture models (MM4, MM5 and MM6) share Y_{nor} , but, in this case, the curves are generated by

$$y(s) = u_1 \sin s + u_2 \cos s, \quad (4.5)$$

where $s \in [0, 2\pi]$ and u_1 and u_2 are observations from a continuous uniform random variable between 0.05 and 0.15.

MM1, MM2 and MM3 differ in their Y_{out} components. Under MM1, the outliers are generated by

$$y(s) = 8s - 2 + \epsilon(s),$$

which produces faint outliers of both shape and low magnitude nature. Under MM2, the outliers are generated by adding to (4.4) an observation from a $N(0, 1)$, and as result outliers are more irregular than normal curves. Finally, under MM3, the outliers are generated by

$$y(s) = 4 \exp(s) + \epsilon(s),$$

which produces curves that are normal in the first part of the domain, but that become exponentially outlying.

Similarly, MM4, MM5 and MM6 differ in their Y_{out} components. Under MM4, the outliers are generated replacing u_2 with u_3 in (4.5), where u_3 is an observation from a continuous uniform random variable between 0.15 and 0.17. This change produces partial low magnitude outliers in the first and middle part of the domain of the curves. Under MM5, the outliers are generated by adding to (4.5) an observation from a $N(0, (\frac{0.1}{2})^2)$, and as result outliers are more irregular curves. Finally, under MM6, the outliers are generated by

$$y(s) = u_1 \sin s + \exp\left(\frac{0.69s}{2\pi}\right) u_4 \cos s, \quad (4.6)$$

where u_4 is an observation from a continuous uniform random variable between 0.1 and 0.15. As MM3, MM6 allows outliers that are normal in the first part of the domain and become outlying with an exponential pattern. In [Figure 4.1](#) we report a simulated dataset with at least one outlier for each mixture model.

The details of the simulation study are the following: for each mixture model, we generate 100 datasets, each one composed of 50 curves. Two values of the contamination probability α are considered: 0.02 and 0.05. All curves are generated using a discretized and finite set of 51 equidistant points in the domain of each mixture model ($[0, 1]$ for MM1, MM2 and MM3; $[0, 2\pi]$

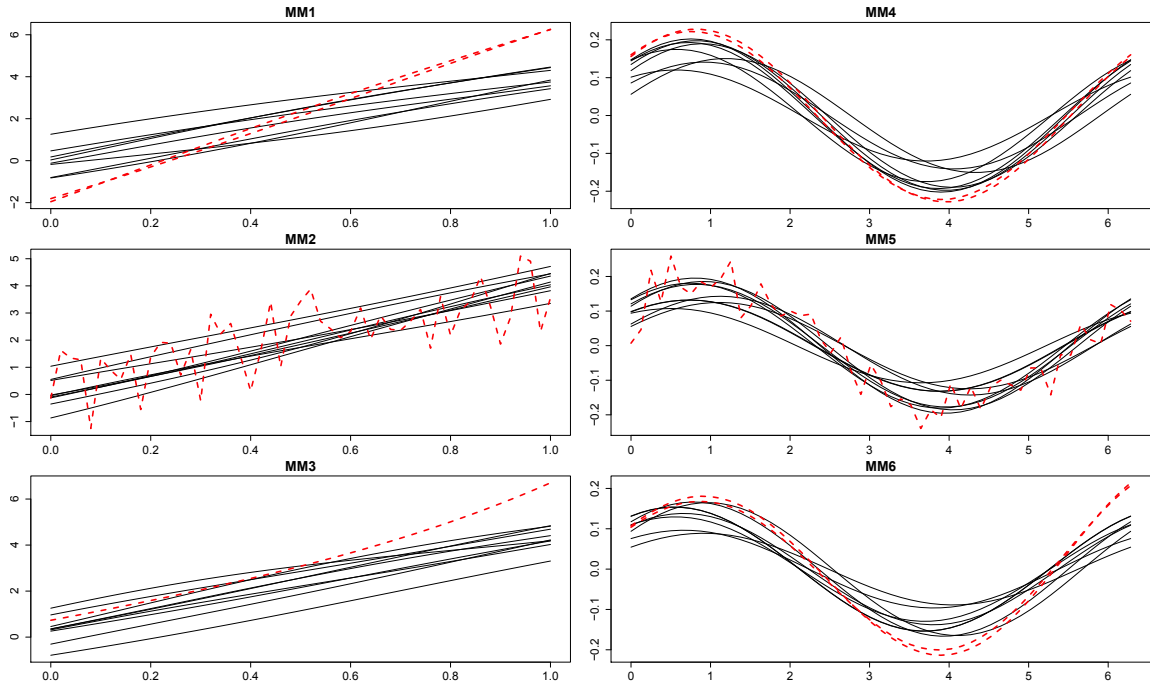


Figure 4.1: Examples of contaminated functional datasets generated by MM1, MM2, MM3, MM4, MM5 and MM6. Solid curves are normal curves and dashed curves are outliers

for MM4, MM5 and MM6) and the discretized versions of the functional depths are used.

In relation with the methods and the functional depths that we consider in the study, the specifications that we use are the following:

1. FBP when used with FMD, HMD, RTD, IDD, MBD, FSD and KFSD: regarding FBP, as reported in [Section 4.2](#), the central region is built considering the 50% deepest curves and the non-outlying region by inflating by 1.5 times the central region. Regarding the depths, for HMD, we follow the recommendations in [Febrero et al. \(2008\)](#), that is, \mathbb{H} is the L^2 space, $\kappa(x, y) = \frac{2}{\sqrt{2\pi}} \exp\left(-\frac{\|x-y\|^2}{2h^2}\right)$ and h is equal to the 15% percentile of the empirical distribution of $\{\|y_i - y_j\|, y_i, y_j \in Y_n\}$. For RTD and IDD, we work with 50 projections in random Gaussian directions. For MBD, we consider bands defined by two curves. For FSD and KFSD, we assume that the curves lie in the L^2 space. Moreover, in KFSD we set σ equal to a moderately local percentile (50%) of the empirical distribution of $\{\|y_i - y_j\|, y_i, y_j \in Y_n\}$.

2. B_{tri} and B_{wei} when used with FMD, HMD, RTD, IDD, MBD, FSD and KFSD: $\gamma = 0.05$, $B = 100$, $\alpha_T = \alpha$. Regarding the depths, we use the specifications reported for FBP.
3. FBG: as reported in [Section 4.2](#), the central region is built considering the 50% deepest bivariate robust functional principal component scores and the non-outlying region by inflating by 2.58 times the central region.
4. FHD: $\beta = \alpha$.
5. $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$: $n_Y = n = 50$ (since $Y_{n_Y} = Y_n$), $\gamma = 0.05$, $\alpha_T = \alpha$ (only for $KFSD_{tri}$), $n_Z = 6n$, $\delta = 0.05$, $r = \alpha$, desired FAP = 0.10. Moreover, as introduced in [Chapter 2](#), we consider a set of percentiles $\{p_k\} = \{p_1, \dots, p_K\}$ to set σ in KFSD. In outlier detection, $\{p_k\} = \{10k\%, k = 1, \dots, K = 9\}$. The way in which we propose to choose the most suitable percentile for outlier detection is presented below.

In supervised classification the availability of training curves with known class memberships makes possible the definition of some natural procedures to set σ for KFSD, such as cross-validation. However, in an outlier detection problem it is common to have no information whether curves are normal or outliers. Therefore, training procedures are not immediately available.

We propose to overcome this drawback by obtaining a “training sample of peripheral curves”, and then choosing the percentile that ranks better the peripheral curves as final percentile for KFSD in $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$. Next, we describe the procedure, which is based on J replications. Let Y_n be the functional dataset on which outlier detection has to be done and let $Y_{(n)} = \{y_{(1)}, \dots, y_{(n)}\}$ be the depth-based ordered version of Y_n , where $y_{(1)}$ and $y_{(n)}$ are the curves with minimum and maximum depth, respectively. The steps to obtain a set of peripheral curves are the following:

1. Let $\{p_1, \dots, p_K\}$ be the set of percentiles in use (in our case, $\{p_k\} = \{10k\%, k = 1, \dots, K = 9\}$), and choose randomly a percentile from the set. For the j th replication, $j \in \{1, \dots, J\}$, denote the selected percentile as p^j . From now on, we use $J = 20$.

2. Using p^j , compute $KFSD_{p^j}(y_i, Y_n)$, $i = 1, \dots, n$, where the notation $KFSD_{p^j}(\cdot, \cdot)$ is used to describe what percentile is used. For the j th replication, denote the KFSD-based ordered curves as $y_{(1),j}, \dots, y_{(n),j}$.
3. Take $y_{(1),j}, \dots, y_{(l_j),j}$, where $l_j \sim \text{Bin}(n, \frac{1}{n})$. Apply the smoothing step described in [Section 4.2](#) to these curves. For the smoothing step, we use Σ_{Y_n} and $\gamma = 0.05$. For the j th replication, denote the peripheral and smoothed curves as $y_{(1),j}^*, \dots, y_{(l_j),j}^*$.
4. Repeat J times steps 1.-3. to obtain a collection of $L = \sum_{j=1}^J l_j$ peripheral curves, say Y_L (for an example, see [Figure 4.2](#)).

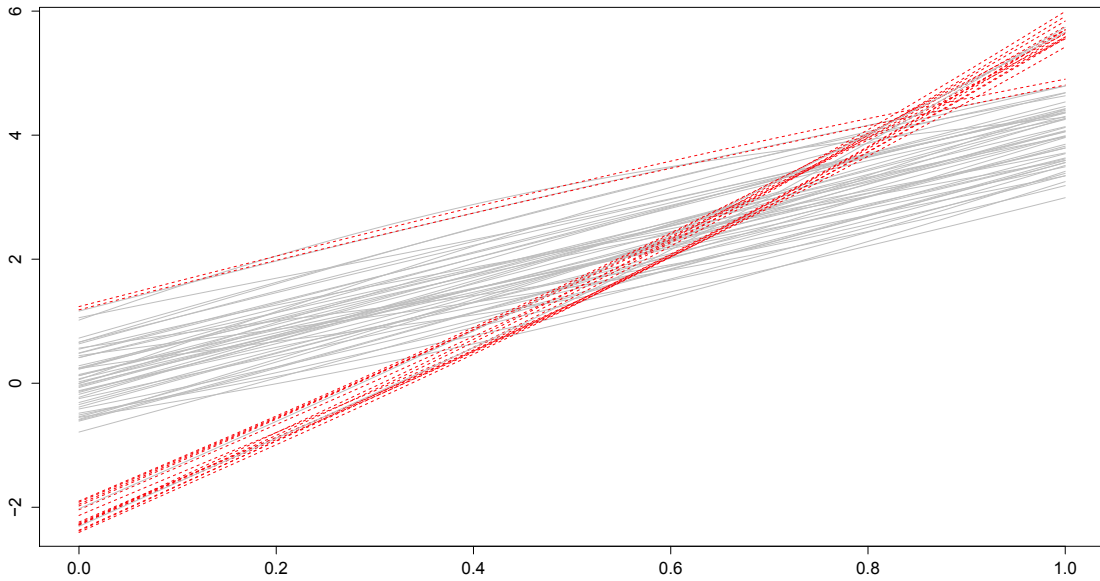


Figure 4.2: Example of a training sample of peripheral curves for a contaminated dataset generated by MM1 with $\alpha = 0.05$. The solid and shaded curves are the original curves (both normal and outliers). The dashed curves are the peripheral curves to use as training sample

Next, Y_L acts as training sample according to the following steps: for each $y_{(i),j}^* \in Y_L$, ($i \leq l_j$), and $p_k \in \{p_1, \dots, p_K\}$, compute $KFSD_{p_k}(y_{(i),j}^*, Y_{-(i),j})$, where $Y_{-(i),j} = Y_n \setminus \{y_{(i),j}\}$. At the end, a $L \times K$ matrix is obtained, say $D_{LK} = \{d_{lk}\}_{\substack{l=1, \dots, L \\ k=1, \dots, K}}$, whose k th column is composed of the KFSD values of the L training peripheral curves when the k th percentile is employed

in KFSD. Let r_{lk} be the rank of d_{lk} in the vector $\{KFSD_{p_k}(y_1, Y_n), \dots, KFSD_{p_k}(y_n, Y_n), d_{lk}\}$, e.g., r_{lk} is equal to 1 ($n + 1$) if d_{lk} is the minimum (maximum) value in the vector. Let R_{LK} be the result of this transformation of D_{LK} , and sum the elements of each column, obtaining a K -dimensional vector, say \mathbf{R}_K . Since the goal is to assign ranks as lower as possible to the peripheral curves, choose the percentile associated to the minimum value of \mathbf{R}_K . When a tie is observed, we break it randomly.

The comparison among methods is performed in terms of both correct and false outlier detection percentages, which are reported in Tables 4.1-4.6. To ease the reading of the tables, for each model and α , we report in bold the 5 best methods in terms of correct outlier detection percentage (c).² For each model, if a method is among the 5 best ones for both contamination probabilities α , we report its label in bold.

Table 4.1: MM1, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	44.34	1.23	43.86	0.73
FBP+HMD	74.53	0.94	72.81	0.61
FBP+RTD	61.32	0.57	63.16	0.31
FBP+IDD	55.66	0.61	61.84	0.34
FBP+MBD	49.06	1.33	50.44	0.69
FBP+FSD	62.26	0.67	61.84	0.40
FBP+KFSD	66.04	0.86	74.12	0.44
B_{tri} +FMD	0.00	0.92	0.00	1.80
B_{tri} +HMD	72.64	1.43	62.28	1.51
B_{tri} +RTD	8.49	0.37	14.47	0.40
B_{tri} +IDD	12.26	0.39	17.11	0.65
B_{tri} +MBD	0.00	0.67	0.00	1.51
B_{tri} +FSD	1.89	0.84	5.70	1.22
B_{tri} +KFSD	70.75	1.57	57.89	1.49
B_{wei} +FMD	0.00	1.23	0.00	1.53
B_{wei} +HMD	71.70	1.16	46.49	0.57
B_{wei} +RTD	10.38	1.25	7.46	1.17
B_{wei} +IDD	14.15	2.29	14.04	2.62
B_{wei} +MBD	0.00	1.14	0.00	1.30
B_{wei} +FSD	1.89	1.33	3.07	1.17
B_{wei} +KFSD	66.04	0.94	57.02	0.52
FBG	100.00	2.27	97.81	2.37
FHD	48.11	1.00	73.68	2.77
$KFSD_{smo}$	89.62	4.50	85.09	2.58
$KFSD_{tri}$	89.62	4.92	92.11	4.40
$KFSD_{wei}$	97.17	9.44	96.93	6.54

Table 4.2: MM2, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	99.09	1.08	96.39	0.84
FBP+HMD	96.36	0.96	96.39	0.88
FBP+RTD	99.09	0.61	94.78	0.25
FBP+IDD	99.09	0.70	95.18	0.38
FBP+MBD	99.09	1.06	96.39	0.82
FBP+FSD	99.09	0.57	94.78	0.36
FBP+KFSD	98.18	0.63	93.98	0.36
B_{tri} +FMD	0.00	0.96	0.00	2.00
B_{tri} +HMD	94.55	1.60	95.18	1.73
B_{tri} +RTD	5.45	0.37	7.63	0.93
B_{tri} +IDD	6.36	0.45	10.04	0.97
B_{tri} +MBD	0.00	1.08	0.40	2.10
B_{tri} +FSD	4.55	1.06	6.02	1.64
B_{tri} +KFSD	99.09	1.60	96.39	1.56
B_{wei} +FMD	0.00	1.41	0.00	1.39
B_{wei} +HMD	94.55	0.94	83.53	0.32
B_{wei} +RTD	7.27	1.51	8.43	1.89
B_{wei} +IDD	8.18	2.49	8.84	2.86
B_{wei} +MBD	0.00	1.29	0.40	1.54
B_{wei} +FSD	6.36	1.43	4.82	1.41
B_{wei} +KFSD	92.73	0.72	81.53	0.51
FBG	8.18	3.07	4.42	2.95
FHD	7.27	1.88	12.45	5.66
$KFSD_{smo}$	100.00	3.91	95.18	2.76
$KFSD_{tri}$	100.00	5.19	97.99	4.84
$KFSD_{wei}$	100.00	9.20	99.60	6.48

²In presence of tie, we look at the false outlier detection percentage (f), preferring the method with lower f.

Table 4.3: MM3, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	65.69	0.92	49.19	0.97
FBP+HMD	89.22	0.57	85.89	0.63
FBP+RTD	86.27	0.45	76.61	0.34
FBP+IDD	79.41	0.51	70.56	0.38
FBP+MBD	74.51	0.88	59.27	0.84
FBP+FSD	79.41	0.51	73.79	0.42
FBP+KFSD	89.22	0.57	83.06	0.59
B_{tri} +FMD	2.94	0.96	4.84	1.24
B_{tri} +HMD	59.80	1.61	55.65	1.64
B_{tri} +RTD	5.88	0.33	4.03	0.40
B_{tri} +IDD	34.31	0.49	23.79	0.76
B_{tri} +MBD	0.98	1.12	3.63	1.49
B_{tri} +FSD	14.71	1.06	17.74	1.41
B_{tri} +KFSD	59.80	1.65	47.98	1.39
B_{wei} +FMD	2.94	1.10	5.24	0.84
B_{wei} +HMD	59.80	1.25	37.90	0.80
B_{wei} +RTD	19.61	0.92	12.90	0.78
B_{wei} +IDD	29.41	2.67	20.97	2.67
B_{wei} +MBD	0.98	1.31	3.23	1.26
B_{wei} +FSD	16.67	1.10	11.29	0.90
B_{wei} +KFSD	55.88	1.12	41.13	0.72
FBG	86.27	2.65	78.63	1.73
FHD	49.02	1.02	65.73	2.88
$KFSD_{smo}$	89.22	3.90	73.79	2.95
$KFSD_{tri}$	90.20	4.63	83.47	4.71
$KFSD_{wei}$	97.06	8.96	90.32	6.50

Table 4.4: MM4, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	1.02	0.00	0.00	0.00
FBP+HMD	6.12	0.00	1.60	0.02
FBP+RTD	0.00	0.00	0.00	0.00
FBP+IDD	0.00	0.00	0.00	0.00
FBP+MBD	0.00	0.00	0.00	0.00
FBP+FSD	0.00	0.00	0.00	0.00
FBP+KFSD	2.04	0.00	0.80	0.00
B_{tri} +FMD	64.29	0.18	46.80	0.15
B_{tri} +HMD	43.88	0.06	20.40	0.21
B_{tri} +RTD	27.55	1.08	14.80	0.80
B_{tri} +IDD	67.35	0.59	47.60	0.46
B_{tri} +MBD	66.33	0.14	43.20	0.06
B_{tri} +FSD	68.37	0.12	46.80	0.13
B_{tri} +KFSD	57.14	0.24	27.20	0.11
B_{wei} +FMD	51.02	0.12	22.40	0.02
B_{wei} +HMD	40.82	0.04	12.00	0.00
B_{wei} +RTD	24.49	0.18	16.00	0.04
B_{wei}+IDD	90.82	2.26	73.60	1.47
B_{wei} +MBD	56.12	0.08	26.40	0.00
B_{wei} +FSD	61.22	0.08	28.00	0.00
B_{wei} +KFSD	56.12	0.12	20.40	0.00
FBG	9.18	0.53	6.80	1.09
FHD	51.02	1.02	37.60	4.34
$KFSD_{smo}$	87.76	2.16	50.00	1.24
$KFSD_{tri}$	91.84	3.00	64.80	2.91
$KFSD_{wei}$	95.92	5.08	62.00	3.35

Table 4.5: MM5, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	55.56	0.00	54.00	0.00
FBP+HMD	66.67	0.00	68.40	0.04
FBP+RTD	57.58	0.00	54.40	0.00
FBP+IDD	52.53	0.00	56.00	0.00
FBP+MBD	55.56	0.00	55.20	0.00
FBP+FSD	55.56	0.00	55.60	0.00
FBP+KFSD	60.61	0.00	59.20	0.00
B_{tri} +FMD	3.03	0.16	2.80	0.36
B_{tri}+HMD	96.97	0.16	89.20	0.17
B_{tri} +RTD	12.12	1.31	18.40	1.37
B_{tri} +IDD	22.22	0.84	29.20	0.63
B_{tri} +MBD	3.03	0.18	3.20	0.32
B_{tri} +FSD	29.29	0.18	29.20	0.29
B_{tri} +KFSD	90.91	0.27	91.20	0.19
B_{wei} +FMD	3.03	0.22	2.40	0.19
B_{wei} +HMD	93.94	0.02	71.20	0.00
B_{wei} +RTD	16.16	0.41	20.00	0.38
B_{wei} +IDD	23.23	3.20	21.60	2.74
B_{wei} +MBD	4.04	0.24	3.60	0.23
B_{wei} +FSD	26.26	0.12	21.60	0.08
B_{wei} +KFSD	88.89	0.12	68.00	0.04
FBG	0.00	1.02	0.40	0.04
FHD	4.04	1.96	12.80	5.64
$KFSD_{smo}$	98.99	1.82	94.00	0.44
$KFSD_{tri}$	98.99	2.61	98.00	2.11
$KFSD_{wei}$	100.00	4.61	98.40	2.11

Table 4.6: MM6, $\alpha = \{0.02, 0.05\}$. Correct (c) and false (f) outlier detection percentages of FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	$\alpha = 0.02$		$\alpha = 0.05$	
	c	f	c	f
FBP+FMD	48.42	0.00	44.19	0.00
FBP+HMD	60.00	0.18	62.92	0.00
FBP+RTD	55.79	0.00	54.68	0.00
FBP+IDD	46.32	0.00	40.07	0.00
FBP+MBD	48.42	0.00	45.69	0.00
FBP+FSD	52.63	0.00	52.43	0.00
FBP+KFSD	57.89	0.00	56.93	0.00
B_{tri} +FMD	30.53	0.16	35.21	0.32
B_{tri} +HMD	67.37	0.24	50.94	0.15
B_{tri} +RTD	22.11	1.06	17.23	0.61
B_{tri} +IDD	32.63	0.57	20.97	0.51
B_{tri} +MBD	28.42	0.24	31.46	0.36
B_{tri} +FSD	50.53	0.20	44.94	0.21
B_{tri} +KFSD	66.32	0.22	48.31	0.13
B_{wei} +FMD	25.26	0.22	18.35	0.06
B_{wei} +HMD	67.37	0.12	38.95	0.00
B_{wei} +RTD	41.05	0.31	34.46	0.19
B_{wei} +IDD	33.68	2.34	23.22	1.75
B_{wei} +MBD	23.16	0.18	17.98	0.15
B_{wei} +FSD	43.16	0.14	29.59	0.11
B_{wei} +KFSD	64.21	0.14	43.45	0.00
FBG	17.89	0.02	14.98	0.06
FHD	52.63	1.02	61.80	2.85
$KFSD_{smo}$	91.58	2.08	71.16	0.95
$KFSD_{tri}$	93.68	2.69	82.02	2.49
$KFSD_{wei}$	96.84	4.69	83.15	2.75

The results in Tables 4.1-4.6 show that:

1. $KFSD_{tri}$ and $KFSD_{wei}$ are always among the 5 best methods, whereas $KFSD_{smo}$ 10 times over 12, but when its performance is not among the 5 best, it is neither extremely far from the fifth method (MM2, $\alpha = 0.02$: 95.18% against 96.39%; MM3, $\alpha = 0.02$: 73.79% against 78.63%). The rest of the methods are among the 5 best procedures at most 5 times over 12 (FBP+HMD).
2. Regarding MM5 and MM6, our procedures are clearly the best options in terms of correct detection (c), and in the following order: $KFSD_{wei}$, $KFSD_{tri}$ and $KFSD_{smo}$. In general, this pattern is observed overall the simulation study. Note that for MM6 and $\alpha = 0.02$ we observe the best relative performances of $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$, i.e., 91.58%, 93.68% and 96.84%, respectively, against 67.37% of the fourth best method (B_{wei} +HMD), that is, we observe differences greater than 20%, and approaching 30% if $KFSD_{wei}$ and B_{wei} +HMD are compared.
3. About MM3, $KFSD_{wei}$ is clearly the best method in terms of correct detection, however at the price of having a greater false detection (f). This is in general the main weak point of $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$. As for correct detection, we observe a overall pattern in our methods in false detection, but in an opposite way, indicating therefore a trade-off between c and f. Relative high false detection percentages are however something expected in $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$ since these methods are based on the definition of a desired false alarm probability, which is equal to 10% in this study. Concerning MM2, we observe similar results to MM3, but in this case the performances of the best competitors for $KFSD_{wei}$, i.e., $KFSD_{smo}$, $KFSD_{tri}$, FBP-based methods and B_{tri} and B_{wei} when used with local depths, are closer to the results of $KFSD_{wei}$.
4. Finally, there are only 3 cases in which a competitor outperforms our best method: for MM1 and both α the best method is FBG, whereas for MM4 and $\alpha = 0.05$ the best method is B_{wei} +IDD. However, both FBG and B_{wei} +IDD do not show behaviors as stable as

KFSD_{smo} , KFSD_{tri} and KFSD_{wei} do. Indeed, they show poor performances under other scenarios, e.g., MM2, MM5 or MM6.

In [Figure 4.3](#) we report a series of boxplots summarizing which percentiles have been selected in the training steps for KFSD_{smo} , KFSD_{tri} and KFSD_{wei} .

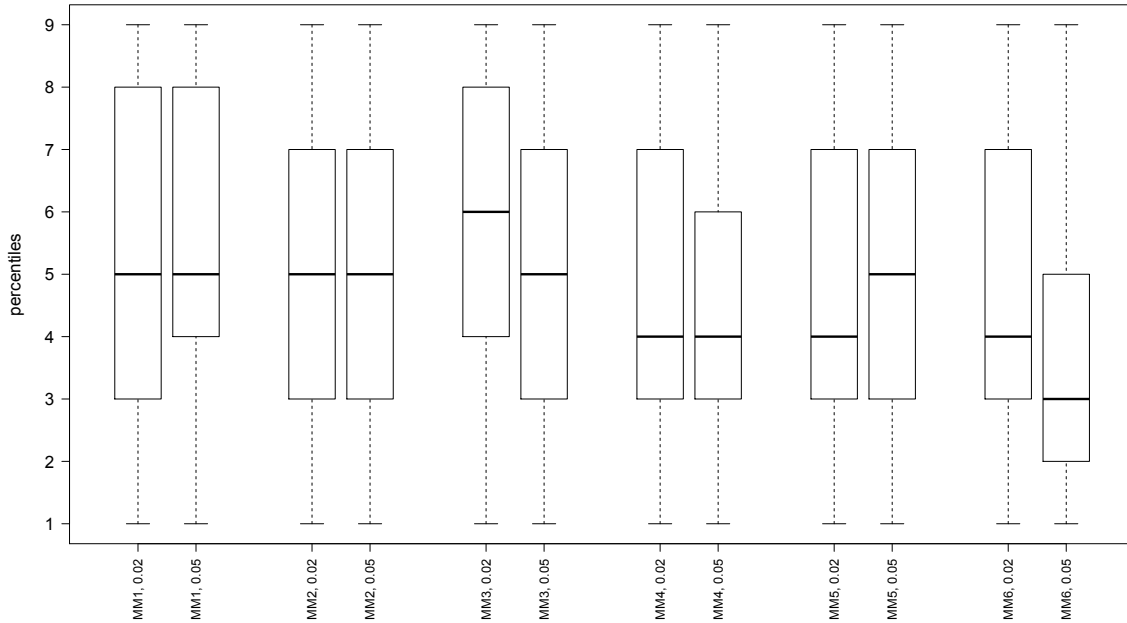


Figure 4.3: Boxplots of the percentiles selected in the training steps of the simulation study for KFSD_{smo} , KFSD_{tri} and KFSD_{wei}

Observing [Figure 4.3](#), the following general remarks can be done. First, MM6 is the mixture model for which lower percentiles have been selected, and it is also the scenario in which our methods considerably outperform their competitors. The need for a more local approach for MM6-data may explain the two observed facts about this mixture model. Second, lower and more local percentiles have been chosen for mixture models with nonlinear mean functions (MM4, MM5 and MM6) than for mixture models with linear mean functions (MM1, MM2 and MM3). Finally, the percentiles selected by means of the proposed training procedure seem to vary among datasets. However, except for MM3 and $\alpha = 0.02$, at least for half of the datasets a percentile not greater than the median has been chosen, which implies at most a moderately local approach.

4.4 Real Data Study: Nitrogen Oxides (NO_x) Data

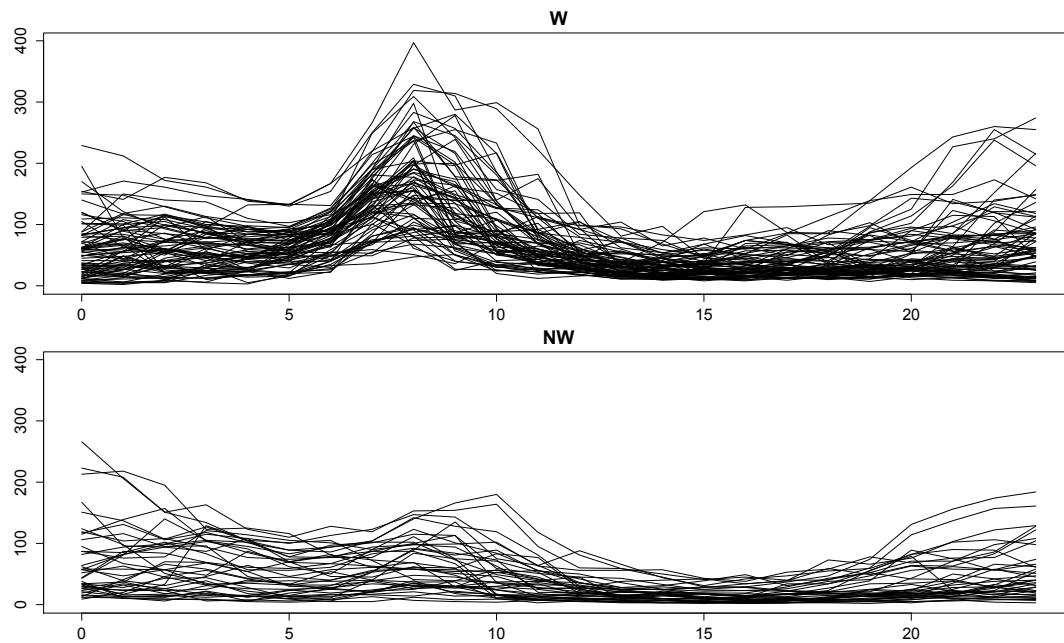
Besides simulated data, we consider a real dataset which consists in nitrogen oxides (NO_x) emission level daily curves measured every hour close to an industrial area in Poblenou (Barcelona). The dataset is available in the R package `fda.usc` (Febrero and Oviedo de la Fuente 2012) and outlier detection on these data has been first performed by Febrero et al. (2008) in the paper where B_{tri} and B_{wei} were presented. We enhance their study by considering more methods and depths.

According to Febrero et al. (2008), NO_x are one of the most important pollutants, and it is important to identify outlying trajectories because these curves may both compromise any statistical analysis and be of special interest for further analysis.

More in details, the NO_x levels were measured in $\mu\text{g}/\text{m}^3$ every hour of every day for the period 23/02/2005-26/06/2005, but only for 115 days was possible to measure the NO_x at every hour. These 115 curves are the ones composing the final NO_x dataset. However, since the NO_x dataset clearly includes working as well as nonworking day curves, following Febrero et al. (2008), it is more appropriate to consider two different datasets, that is, a sample of 76 working day curves (from now on, W) and another of 39 nonworking day curves (from now on, NW). Both W and NW are showed in Figure 4.4: at first glance, it seems that each dataset may contain outliers, especially partial outliers.

Therefore, because of the possible presence of faint outliers, a local depth approach by means of the use of KFSD_{smo} , KFSD_{tri} and KFSD_{wei} may be a good strategy to detect outliers. Besides them, we do outlier detection with all the methods used in Section 4.3. For all the procedures we use the same specifications as in Section 4.3, and we assume $\alpha = 0.05$. For each method, we report the labels of the curves detected as outliers in Table 4.7, whereas in Figure 4.5 we highlight these curves.

For what concerns W, most of the methods detect as outlier day 37, which apparently shows a partial outlying behavior before noon and at the end of the day. Another day detected as outlier by many methods is day 16, whose curve is the one with the highest morning peak. In addition to curves 16 and 37, KFSD_{smo} detects as outlier curve 14, as other nine methods

Figure 4.4: NO_x data: working (top) and non working (bottom) day curves.Table 4.7: NO_x data, W and NW datasets. Curves detected as outliers by FBP, B_{tri} , B_{wei} , FBG, FHD, $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$

	w days	non w days
	detected outliers	
FBP+FMD	-	-
FBP+HMD	12, 16, 37	5, 7, 20, 21
FBP+RTD	37	20
FBP+IDD	-	5, 7, 20
FBP+MBD	-	-
FBP+FSD	37	-
FBP+KFSD	12, 16, 37	5, 7, 20, 21
B_{tri} +FMD	16, 37	7
B_{tri} +HMD	14, 16, 37	7, 20
B_{tri} +RTD	14, 16, 37	-
B_{tri} +IDD	11, 14, 16, 37	-
B_{tri} +MBD	16, 37	7
B_{tri} +FSD	14, 16, 37	-
B_{tri} +KFSD	12, 14, 16, 37	7, 20
B_{wei} +FMD	16	7
B_{wei} +HMD	16, 37	7, 20
B_{wei} +RTD	14, 16, 37, 38	-
B_{wei} +IDD	16, 37	-
B_{wei} +MBD	16	7
B_{wei} +FSD	16, 37	-
B_{wei} +KFSD	16, 37	7, 20
FBG	16, 37	-
FHD	12, 14, 16, 37	7, 20
$KFSD_{smo}$	14, 16, 37	7, 20, 21
$KFSD_{tri}$	12, 14, 16, 37	7, 20, 21
$KFSD_{wei}$	11, 12, 13, 14, 15, 16, 37, 38	7, 20, 21

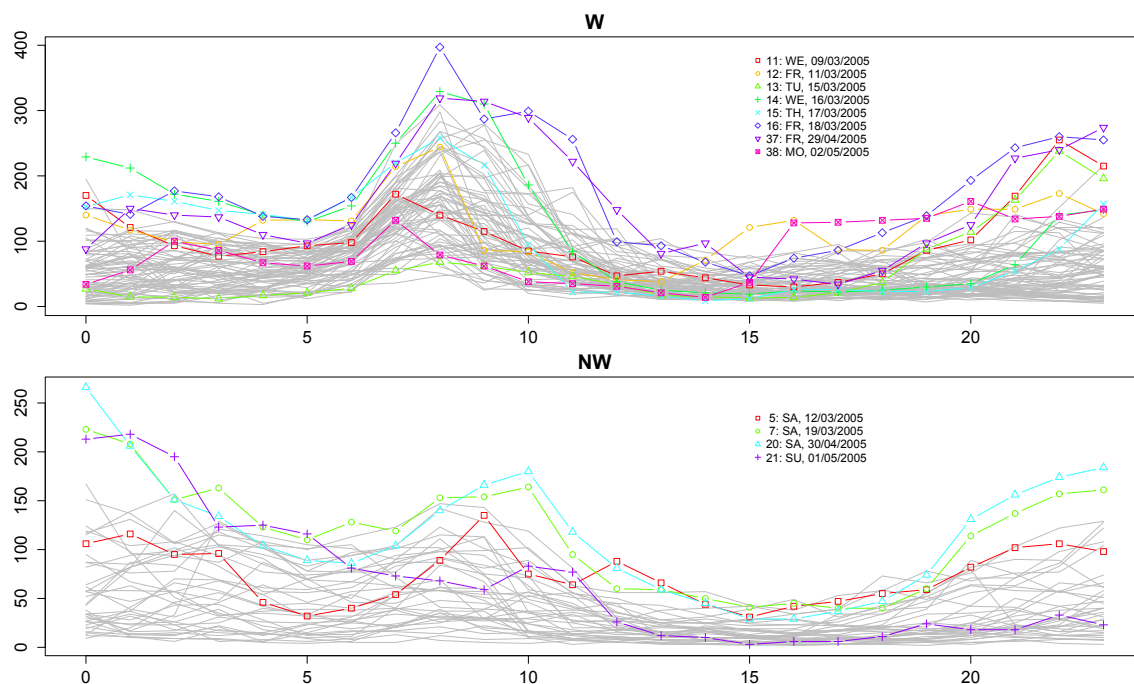


Figure 4.5: NO_x dataset, curves detected as outliers in Table 4.7: working (top) and non working (bottom) days

do, recognizing a seeming outlying pattern in early hours of the day. Additionally, KFSD_{tri} includes among the outliers also day 12, which may be atypical because of its behavior in early afternoon. Finally, KFSD_{wei} detects as outliers the greatest number of curves. This last result may appear exaggerated, but all the curves that are outliers according to KFSD_{wei} seem to have some partial deviations from the majority of curves. For example, day 13, whose curve is considered normal by the rest of the procedures, shows a peak at end of the day. Similar peaks can be observed also in other curves detected as outliers by other methods (e.g., days 16 and 37), which means that it may be occurring a masking effect to day 13's detriment, and only KFSD_{wei} points out this possibly outlying feature of the curve. Regarding the training step for KFSD to set σ , it gives as result the 70% percentile. Observing the first graph of Figure 4.4, it can be noticed that some curves have a likely outlying behavior, and this may be the reason why a weakly local approach for KFSD may be adequate enough.

In the case of NW, some methods detect no curves as outliers (e.g., all the global FSD-based methods), exclusively three FBP-based methods flag day 5 as outlier, whereas days 7, 20

and 21 are detected as outliers by, among others, our methods. Days 7 and 20, which have two peaks, at the beginning and end of the day, are also flagged by other twelve and eight methods, respectively, while day 21, which shows a single peak in the first hours of the day, is considered atypical by only two other methods, which happen to be local (FBP+HMD and FBP+KFSD). This last result may be connected with what has been observed at the KFSD training step for selecting the percentile, i.e., the selection of the 30% percentile. Therefore, KFSD_{smo} , KFSD_{tri} and KFSD_{wei} work with a strongly local percentile, and their results partially resemble the ones of the previously mentioned local techniques.

4.5 Conclusions

This chapter proposes three methods to detect outliers in functional samples based on KFSD. We presented a way to set a KFSD-threshold to identify outliers in [Theorem 4.1](#). In practice, it is necessary to observe two samples to apply [Theorem 4.1](#), and one sample is requested to have a considerably large size. To overcome this practical limitation, we proposed KFSD_{smo} , KFSD_{tri} and KFSD_{wei} : they are methods based on smoothed resampling techniques and, more important, can be applied when a unique functional sample is available, no matter its size.

We also proposed a new procedure to set the key parameter of KFSD, i.e., its bandwidth σ , based on obtaining training samples by means of smoothed resampling techniques. The general idea behind this procedure can be applied to other functional depths or methods with parameters that need to be set.

We investigated the performances of KFSD_{smo} , KFSD_{tri} and KFSD_{wei} by means of a simulation study. We focused on challenging scenarios with low magnitude, shape and partial outliers (faint outliers) instead of high magnitude outliers (clear outliers). Along the simulation study, KFSD_{wei} , KFSD_{tri} and KFSD_{smo} attained the largest correct detection performances in most of the analyzed setups. However, in some cases they paid a price in terms of false detection. Nevertheless, KFSD_{wei} , KFSD_{tri} and KFSD_{smo} work with a given desired false alarm probability. Thus, higher false detection percentages than the competitors can be explained

by the inherent structure of the methods. Concerning the remaining methods, few competitors in few scenarios outperformed our methods. However, in these cases the differences were not great, especially for KFSD_{wei} and KFSD_{tri} , and more important, these competitors did not show stability across scenarios in their results. Finally, we also considered a real application consisting in NO_x emission daily curves.

4.6 Appendix

As explained in Section 4.2, Theorem 4.1 is a functional extension of a result derived by Chen et al. (2009) for KSD, and since they are closely related, next we report a sketch of the proof of Theorem 4.1. The proof for KSD is mostly based on an inequality known as McDiarmid's inequality (McDiarmid 1989), which also applies to general probability spaces, and therefore to functional Hilbert spaces. We report this inequality in the next lemma:

Lemma 4.1. (McDiarmid 1989 [1.2]) *Let $\Omega_1, \dots, \Omega_n$ be probability spaces. Let $\Omega = \prod_{j=1}^n \Omega_j$ and let $X : \Omega \rightarrow \mathbb{R}$ be a random variable. For any $j \in \{1, \dots, n\}$, let $(\omega_1, \dots, \omega_j, \dots, \omega_n)$ and $(\omega_1, \dots, \hat{\omega}_j, \dots, \omega_n)$ be two elements of Ω that differ only in their j th coordinates. Assume that X is uniformly difference-bounded by $\{c_j\}$, that is, for any $j \in \{1, \dots, n\}$,*

$$|X(\omega_1, \dots, \omega_j, \dots, \omega_n) - X(\omega_1, \dots, \hat{\omega}_j, \dots, \omega_n)| \leq c_j. \quad (4.7)$$

Then, if $\mathbb{E}[X]$ exists, for any $\tau > 0$

$$\Pr(X - \mathbb{E}[X] \geq \tau) \leq \exp\left(\frac{-2\tau^2}{\sum_{j=1}^n c_j^2}\right).$$

In order to apply Lemma 4.1 to our problem, define

$$X(z_1, \dots, z_{n_Z}) = -\frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y} | Y_{n_Y}), \quad (4.8)$$

whose expected value is given by

$$\mathbb{E}[X] = \mathbb{E} \left[-\frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y} | Y_{n_Y}) \right] = -\mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y} | Y_{n_Y})]. \quad (4.9)$$

Now, for any $j \in \{1, \dots, n_Z\}$ and $\hat{z}_j \in \mathbb{H}$, the following inequality holds

$$|X(z_1, \dots, z_j, \dots, z_{n_Z}) - X(z_1, \dots, \hat{z}_j, \dots, z_{n_Z})| \leq \frac{1}{n_Z},$$

and it provides assumption (4.7) of Lemma 4.1. Therefore, for any $\tau > 0$

$$\Pr \left(\mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y} | Y_{n_Y})] - \frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y} | Y_{n_Y}) \geq \tau \right) \leq \exp(-2n_Z \tau^2),$$

and by the law of total probability

$$\begin{aligned} & \mathbb{E} \left[\Pr \left(\mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y} | Y_{n_Y})] - \frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y} | Y_{n_Y}) \geq \tau \right) \right] \\ &= \Pr \left(\mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y})] - \frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y}) \geq \tau \right) \leq \exp(-2n_Z \tau^2) \end{aligned}$$

Next, setting $\delta = \exp(-2n_Z \tau^2)$, and solving for τ , the following result is obtained:

$$\tau = \sqrt{\frac{\ln 1/\delta}{2n_Z}}.$$

Therefore,

$$\Pr \left(\mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y})] \leq \frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y}) + \sqrt{\frac{\ln 1/\delta}{2n_Z}} \right) \geq 1 - \delta. \quad (4.10)$$

However, Theorem 4.1 provides a probabilistic upper bound for $\mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})]$. First, note that

$$\mathbb{E}_{x \sim Y_{mix}} [g(x, Y_{n_Y})] = (1 - \alpha) \mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})] + \alpha \mathbb{E}_{x \sim Y_{out}} [g(x, Y_{n_Y})],$$

and then, for $\alpha > 0$,

$$\mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})] \leq \frac{1}{1 - \alpha} \mathbb{E}_{x \sim Y_{mix}} [g(x, Y_{n_Y})] = \frac{1}{1 - \alpha} \mathbb{E}_{z_1 \sim Y_{mix}} [g(z_1, Y_{n_Y})]. \quad (4.11)$$

Consequently, combining (4.10) and (4.11), we obtain

$$\Pr \left(\mathbb{E}_{x \sim Y_{nor}} [g(x, Y_{n_Y})] \leq \frac{1}{1 - r} \left[\frac{1}{n_Z} \sum_{i=1}^{n_Z} g(z_i, Y_{n_Y}) + \sqrt{\frac{\ln 1/\delta}{2n_Z}} \right] \right) \geq 1 - \delta,$$

which completes the proof. ■

*Marta: Quando c'è l'amore c'è tutto.
Gaetano: No, chell' è 'a salute!
(In Ricomincio da tre)*

Chapter 5

Conclusions

This dissertation was framed in the statistical field of functional data analysis (FDA). In particular, our main interest was the notion of functional depth. A functional depth allows to obtain a center-outward ordering criterion for a functional sample Y_n and to assess whether a curve is central or peripheral relative to the rest of the curves in the sample. Besides ranking curves, we showed that a functional depth may be useful to estimate in a robust way the center of a functional random variable Y , to obtain central regions or to build other robust procedures based on the identification of the most representative/central curves and on the elimination or down-weighting of the least representative/central functional observations. For these reasons, in the last 15 years several authors have proposed different functional depths. We reported some of them in [Chapter 2](#).

One of these proposals is the functional spatial depth (FSD, [Chakraborty and Chaudhuri 2014](#)). FSD relies on both a spatial and global approach. The concept of global approach of a depth can be described as follows: let $y_i, y_j, y_k \in Y_n$ be two central and one peripheral curve relative to Y_n , respectively. Then, according to a global depth, the contributions of y_j and y_k to the depth value of y_i should be the same. However, an alternative and also reasonable approach may consist in reducing the contribution of y_k to the depth value of y_i since they are two distant curves. We described this different approach as local. The main contribution of this thesis was the definition of a new functional depth that is precisely based on a local approach,

that is, the kernelized functional spatial depth (KFSD). Indeed, KFSD is a local-oriented version of FSD and the transition from global (FSD) to local (KFSD) was achieved using kernel methods. Moreover, we provided also another way to interpret the connection between FSD and KFSD: using an embedding map ϕ from an infinite-dimensional Hilbert space \mathbb{H} onto a feature space \mathbb{F} , it is possible to see KFSD as a recoded version of FSD. The double nature of KFSD as a kernel-based and recoded depth is explained by the existing link between kernel functions and embedding maps through the inner product function defined on \mathbb{H} .

In [Chapter 3](#) KFSD was used together with three different supervised classification methods (DTM, WAD and WMD) to assign unlabeled test curves to one of the groups of some labeled training curves. We considered both simulated and real curves and DTM, WAD and WMD were used in conjunction with KFSD and other functional depths (FMD, HMD, RTD, IDD, MBD and FSD). The functional extension of k -NN was considered as benchmark procedure. Therefore, we analyzed a total of 22 classification methods and we substantially contributed to a better knowledge of supervised functional classification. Moreover, we designed a rather original structure for our study since we focused on challenging classification problems where the different classes of curves were difficult to recognize at first sight and/or were contaminated with outliers. According to the results reported in [Chapter 3](#), KFSD-based methods behave well in such difficult classification problems, and in particular the pair WMD+KFSD resulted as the best and most stable method.

Finally, in [Chapter 4](#) we dealt with another functional problem, that is, outlier detection. The reason why it is a good idea to use functional depths to recognize atypical curves is the following: if any outlier is present in a functional sample, a functional depth is expected to assign a low depth value to this atypical curve. However, even if it is natural to build outlier detection methods based on the notion of functional depth, low depth values may also be assigned to peripheral but non-outlying curves. Therefore, it is necessary to define procedures able to provide a threshold depth value to discriminate between normal curves and outliers. Indeed, in [Chapter 4](#) we proposed three different methods to select a threshold for KFSD (KFSD_{smor} , KFSD_{tri} and KFSD_{wei}) and we compared these new proposals with other depth-based (FBP,

B_{tri} and B_{wei}) and non-depth-based (FBG and FHD) existing procedures. As for classification, we developed an extensive simulation study. In this case we allowed exclusively faint outliers, which are outliers hard to be recognized, and we overlooked clear outliers, which are instead easily identifiable by most outlier detection procedures. The results of the simulation study showed that the new methods $KFSD_{smo}$, $KFSD_{tri}$ and $KFSD_{wei}$ based on KFSD had the best and most stable performances. This was especially the case of $KFSD_{wei}$ and $KFSD_{tri}$.

5.1 Research Lines

We close this dissertation describing some aspects that may characterize our future research:

1. Properties of KFSD

In [Section 2.4](#) we presented KFSD which can be interpreted as a local-oriented version of the global oriented depth FSD. Moreover, we also showed that KFSD and FSD are related since the first can be interpreted as a recoded version of the second. Along the dissertation this connection remained rather implicit because we exploited another important relationship, that is, the link between an embedding map ϕ and a kernel function κ . However, a profounder study of the relationship between KFSD and FSD through ϕ may result useful for the definition of some properties of KFSD. Indeed, as reported in [Section 2.3](#), [Chakraborty and Chaudhuri \(2014\)](#) stated some properties of FSD: their extension to KFSD is one of our main future research goals. Also, the fulfillment by KFSD of the desirable properties for a functional depth presented by [Mosler and Polyakova \(2012\)](#) is among our objectives.

2. Different kernel functions for KFSD

At this stage of our research we did not investigate different kernels functions than $\kappa(x, y) = \exp\left(-\frac{\|x-y\|^2}{\sigma^2}\right)$ for KFSD. However, while it is rather common to use Gaussian kernel functions in kernel-based methods, it would be interesting to explore how the choice of different kernels would affect the behavior of KFSD in supervised classi-

fication, outlier detection or other problems. For example, besides [Cuevas et al. \(2006\)](#), also [Dabo-Niang et al. \(2007\)](#) proposed a kernel-type estimator of the modal curve. They used a kernel function of a different nature, that is,

$$\kappa(x, y) = \frac{3}{2} \left(1 - \frac{d(x, y)^2}{h^2} \right) 1_{(0,1)}(d(x, y)),$$

where $d(x, y)$ is some measure of proximity between x and y , h is a bandwidth and $1_{(0,1)}(u)$ is the indicator function of the interval $(0, 1)$ of $d(x, y)$. Also, the review on the classes of kernels for machine learning from a statistics perspective by [Genton \(2002\)](#) may offer other interesting alternatives for the $\kappa(x, y)$ in KFSD. Therefore, an extensive study that compares the behavior of KFSD with different kernel functions, as well as the definition of decision rules to make the final decision, may further improve the already good performances of KFSD in supervised classification and outlier detection.

3. Issues about KFSD as a descriptive FDA tool

We provided two different procedures to set the bandwidth σ of KFSD in supervised classification and outlier detection. However, KFSD can also be used as a tool to obtain some descriptive statistics of a functional sample Y_n . In this work we did not focus on descriptive FDA and neither provided a rule to set the bandwidth of KFSD in descriptive analyses. Indeed, for such analyses it is not even obvious what criterion to consider to choose σ , unlike in supervised classification and outlier detection. For these reasons, we intend to explore different criteria and decision rules to select the bandwidth of KFSD in descriptive FDA problems.

4. Other KFSD-based methods

As described in the introduction of this thesis, besides supervised classification and outlier detection, there are other problems that can be tackled using the notion of functional depth, and therefore KFSD. For instance, we did not study the performances of KFSD in location estimation of the center of a functional random variable Y and such study would enhance our knowledge about the usefulness of KFSD in FDA. Another natural

step ahead in our research may be the definition of KFSD-based cluster analysis procedures. Note that outlier detection, that has been studied in this thesis, can be seen as a special case of cluster analysis since it is a cluster problem with maximum two clusters, one of them with size much smaller than the other (even 0). Therefore, our future efforts may go in this direction to define KFSD-based cluster procedures for functional data which may be compared with some non-depth-based cluster techniques, e.g., the method based on the notion of impartial trimming proposed by [Cuesta-Albertos and Fraiman \(2007\)](#) or the Funclust procedure defined by [Jacques and Preda \(2013\)](#).

References

- Agostinelli, C. and Romanazzi, M. (2011). Local depth. *Journal of Statistical Planning and Inference*, 141:817–830.
- Barnett, V. and Lewis, T. (1994). *Outliers in Statistical Data*, volume 3. Wiley New York.
- Biau, G., Bunea, F., and Wegkamp, M. (2005). Functional classification in Hilbert spaces. *IEEE Transactions on Information Theory*, 51:2163–2172.
- Brown, B. M. (1983). Statistical uses of the spatial median. *Journal of the Royal Statistical Society, Series B*, 45:25–30.
- Burba, F., Ferraty, F., and Vieu, P. (2009). k -nearest neighbour method in functional nonparametric regression. *Journal of Nonparametric Statistics*, 21:453–469.
- Cardot, H., Cénac, P., and Zitt, P.-A. (2013). Efficient and fast estimation of the geometric median in Hilbert spaces with an averaged stochastic gradient algorithm. *Bernoulli*, 19:18–43.
- Cérou, F. and Guyader, A. (2006). Nearest neighbor classification in infinite dimension. *ESAIM. Probability and Statistics*, 10:340–355.
- Chakraborty, A. and Chaudhuri, P. (2014). On data depth in infinite dimensional spaces. *Annals of the Institute of Statistical Mathematics*, 66:303–324.
- Chaudhuri, P. (1996). On a geometric notion of quantiles for multivariate data. *Journal of the American Statistical Association*, 91:862–872.

- Chen, Y., Dang, X., Peng, H., and Bart, H. L. (2009). Outlier detection with the kernelized spatial depth function. *IEEE Transactions on Pattern Analysis and Machine Intelligence*, 31:288–305.
- Cuesta-Albertos, J. A. and Fraiman, R. (2007). Impartial trimmed k -means for functional data. *Computational statistics & data analysis*, 51:4864–4877.
- Cuesta-Albertos, J. A. and Nieto-Reyes, A. (2008). The random Tukey depth. *Computational Statistics and Data Analysis*, 52:4979–4988.
- Cuevas, A. (2014). A partial overview of the theory of statistics with functional data. *Journal of Statistical Planning and Inference*, 147:1–23.
- Cuevas, A., Febrero, M., and Fraiman, R. (2006). On the use of the bootstrap for estimating functions with functional data. *Computational Statistics and Data Analysis*, 51:1063–1074.
- Cuevas, A., Febrero, M., and Fraiman, R. (2007). Robust estimation and classification for functional data via projection-based depth notions. *Computational Statistics*, 22:481–496.
- Cuevas, A. and Fraiman, R. (2009). On depth measures and Dual statistics. a methodology for dealing with general data. *Journal of Multivariate Analysis*, 100:753–766.
- Dabo-Niang, S., Ferraty, F., and Vieu, P. (2007). On the using of modal curves for radar waveforms classification. *Computational Statistics & Data Analysis*, 51(10):4878–4890.
- Epifanio, I. (2008). Shape descriptors for classification of functional data. *Technometrics*, 50:284–294.
- Febrero, M., Galeano, P., and González-Manteiga, W. (2008). Outlier detection in functional data by depth measures, with application to identify abnormal nox levels. *Environmetrics*, 19:331–345.
- Febrero, M., Galeano, P., and Gonzalez-Manteiga, W. (2007). A functional analysis of nox levels: location and scale estimation and outlier detection. *Computational Statistics*, 22:411–427.

- Febrero, M. and Oviedo de la Fuente, M. (2012). Statistical computing in functional data analysis: the R package *fda.usc*. *Journal of Statistical Software*, 51:1–28.
- Ferraty, F. and Vieu, P. (2003). Curves discrimination: a nonparametric functional approach. *Computational Statistics and Data Analysis*, 44:161–173.
- Ferraty, F. and Vieu, P. (2006). *Nonparametric Functional Data Analysis : Theory and Practice*. Springer, New York.
- Fraiman, R. and Muniz, G. (2001). Trimmed means for functional data. *Test*, 10:419–440.
- Galeano, P., Joseph, E., and Lillo, R. E. (2014). The Mahalanobis distance for functional data with applications to classification. *Technometrics*, forthcoming. DOI:10.1080/00401706.2014.902774.
- Genton, M. G. (2002). Classes of kernels for machine learning: a statistics perspective. *The Journal of Machine Learning Research*, 2:299–312.
- Hall, P., Poskitt, D., and Presnell, B. (2001). A functional data-analytic approach to signal discrimination. *Technometrics*, 43:1–9.
- Hastie, T., Buja, A., and Tibshirani, R. (1995). Penalized discriminant analysis. *The Annals of Statistics*, 23:73–102.
- Horváth, L. and Kokoszka, P. (2012). *Inference for Functional Data With Applications*. Springer, New York.
- Hyndman, R. J. and Shang, H. L. (2010). Rainbow plots, bagplots, and boxplots for functional data. *Journal of Computational and Graphical Statistics*, 19:29–45.
- Jacques, J. and Preda, C. (2013). Funclust: a curves clustering method using functional random variables density approximation. *Neurocomputing*, 112:164–171.
- James, G. and Hastie, T. (2001). Functional linear discriminant analysis for irregularly sampled curves. *Journal of the Royal Statistical Society, Series B*, 63:533–550.

- Kudraszow, N. L. and Vieu, P. (2013). Uniform consistency of knn regressors for functional variables. *Statistics and Probability Letters*, 83:1863–1870.
- Liu, R. Y. (1990). On a notion of data depth based on random simplices. *Annals of Statistics*, 18:405–414.
- Liu, R. Y., Parelius, J. M., Singh, K., et al. (1999). Multivariate analysis by data depth: descriptive statistics, graphics and inference, (with discussion and a rejoinder by Liu and Singh). *The Annals of Statistics*, 27:783–858.
- López-Pintado, S. and Romo, J. (2006). Depth-based classification for functional data. In Liu, R., Serfling, R., and Souvaine, D. L., editors, *Data Depth: Robust Multivariate Analysis, Computational Geometry and Applications*, DIMACS Series, pages 103–120, Providence. American Mathematical Society.
- López-Pintado, S. and Romo, J. (2009). On the concept of depth for functional data. *Journal of the American Statistical Association*, 104:718–734.
- Marx, B. and Eilers, P. (1999). Generalized linear regression on sampled signals and curves: a p-spline approach. *Technometrics*, 41:1–13.
- McDiarmid, C. (1989). On the method of bounded differences. In *Survey in Combinatorics*, pages 148–188. Cambridge University Press, Cambridge.
- Mosler, K. and Polyakova, Y. (2012). General notions of depth for functional data. *arXiv preprint*. arXiv:1208.1981.
- Ramsay, J. O. and Silverman, B. W. (2005). *Functional Data Analysis*. Springer, New York.
- Serfling, R. (2002). A depth function and a scale curve based on spatial quantiles. In Dodge, Y., editor, *Statistical Data Analysis Based on the L1-Norm and Related Methods*, pages 25–38. Birkhäuser, Basel.
- Serfling, R. (2006). Depth functions in nonparametric multivariate inference. In Liu, R., Serfling, R., and Souvaine, D. L., editors, *Data Depth: Robust Multivariate Analysis, Computational*

Geometry and Applications, DIMACS Series, pages 1–16, Providence. American Mathematical Society.

Sguera, C., Galeano, P., and Lillo, R. (2014a). Functional outlier detection with a local spatial depth. *Working Paper UC3M*, 14-14 (10).

Sguera, C., Galeano, P., and Lillo, R. (2014b). Spatial depth-based classification for functional data. *TEST*, forthcoming. DOI:10.1007/s11749-014-0379-1.

Sun, Y. and Genton, M. G. (2011). Functional boxplots. *Journal of Computational and Graphical Statistics*, 20:316–334.

Sun, Y., Genton, M. G., and Nychka, D. W. (2012). Exact fast computation of band depth for large functional datasets: How quickly can one million curves be ranked? *Stat*, 1:68–74.

Tukey, J. W. (1975). Mathematics and the picturing of data. In *Proceedings of the International Congress of Mathematicians*, volume 2, pages 523–531.

Zuo, Y. and Serfling, R. (2000). General notions of statistical depth function. *Annals of Statistics*, 28:461–482.