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Changing Dynamics of Foreign Direct Investment in China's Automotive Industry

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Abstract

China's automotive industry has developed dramatically in recent years as more and more major multinational corporations (MNCs) in this industry began to invest in China. Most of these investments have developed in the form of joint-ventures with Chinese state owned enterprises (SOEs). This paper contributes to the current literature by studying the effect of foreign direct investment (FDI) on the productivity of the automotive industry in China using panel data during the 1999–2008 period. Channels through which FDI may directly and indirectly affect the productivity are investigated using pooled ordinary least squares model (POLS) and fixed effects model (FES) to estimate the influence of FDI on productivity in the automotive industry. The results suggest that FDI plays a negative role in this industry and suggests that there is a need for Chinese government to modify its policies and practices in order to improve the productivity of such a key industry in the Chinese economy.

Keywords: FDI, Automotive Industry, Productivity Spillovers, China, Emerging Markets



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I. Introduction

Automobile industry has been the main driver of the intensification of technological changes in the 19th century (Womack, Jones, and Roos, 1990). More importantly, however, in recent years, automobile industry has been one of the most important heritors of Foreign Direct Investment (FDI), especially in emerging markets. The importance of automotive industry is very well accepted in the field of international

business as it contributes to the economic development of any region where it is established. This is mostly due the fact that when established it creates millions of direct and indirect manufacturing employment, and hence generates growth of related upstream and downstream industries. In the United States, for example, the automotive industry and its related industries comprise 10 % of the GDP (Maxton and Wormald, 2004). In the developing countries, a burgeoning domestic auto industry is a key contributing factor of the industrialization process. This is especially true in the case of China.

However, industrial development is not a new phenomenon in China. The Chinese auto industry developed rapidly after economic reform and a policy of openness to business were implemented with the open door policy since 1978.

Yet, despite the economic reforms and increased openness, a large quantity of automobiles was still imported to satisfy the domestic market demand. In the beginning, FDI entered into China through joint ventures and the first joint venture in China's automotive industry was established between the Shanghai Auto Factory and the German Volkswagen in 1985. Since then, several multinational corporations (MNCs) have invested in the Chinese automotive industry. Joint venture operations continued in the 1990s and major automotive industry MNCs cooperated with their Chinese partners to establish joint-ventures. After China's admittance into the World Trade Organization (WTO) in December 2001, the domestic auto production increased dramatically. In 2009, China produced more than 13 million vehicles, which was equivalent to 18 % of the total world production,

and thus became the largest automotive producer surpassing the US and Japan (Chang, 2010).

According to previous literature, FDI plays an important role in the development of China's automotive industry. In theory, FDI promotes the host country's industrial productivity through the following: 1) the development of new products and processes; 2) the demonstration-imitation effect; and 3) the linkages effect and the worker training effect (Romer, 1990; Grossman and Helpman, 1991; Markusen and Venables, 1999).

However, previous literature have also suggested that at times the industrial productivity in a host country may not benefit from FDI because of technology diffusion restrictions imposed by MNCs, particularly those with affiliations in the host countries that decrease the linkage effects or keep the

skills and the know-how secret (Teece, 1977; Das, 1987; Caves, 1996). To better understand the contradictory results that previous literature offers and the relationship between FDI and productivity of the Chinese automotive industry, empirical analysis is required.

Hence, the purpose of our empirical investigation is to estimate the effects of FDI on the productivity of the Chinese automotive industry during the period of 1999–2008. Specifically, we examine the channels through which FDI may affect the productivity of the auto industry and whether the interaction between FDI and human capital can influence the FDI–productivity link. The paper is organized as follows: in Section II, a literature review is presented. In Section III, the model, data and methodology are described. In Section IV, the results are discussed and in

Section V, the conclusion, the limitations of the present research as well as recommendations for further research are presented.

II. Literature Review

According to the surveyed literature of theories on the FDI–productivity links, there are five interrelated modes through which FDI may impact a host country’s productivity directly and indirectly (Caves, 1996; Markusen and Venables, 1999). The direct effect of FDI is defined as the impact on the productivity of firms that results from receiving FDI. The introduction of capital, new products, ideas and practices, new management skills lead to direct transfers of technology. The establishment of R&D centers is also considered a direct effect of FDI.

The indirect effect of FDI, however, is the influence that a MNC's presence has on the productivity of local firms in the form of spillovers from foreign firms to local ones. In other words, what MNCs attempt to keep as proprietary knowledge and technology, will eventually result in indirect transfers of technology (Blomström and Persson, 1994). For example, backward and forward linkages, training effects, demonstration-imitation effects and competition effects are observed in those spillovers.

Direct Effects of FDI

New ideas, products and procedures: Here, new technologies can be introduced with the presence of FDI in the form of new ideas, products and procedures. New skills to operate the technologies are introduced and developed by FDI (Das, 1987; Grossman

and Helpman, 1991). Furthermore, a host country's stock of ideas can be augmented by those new ideas brought by MNCs, thus innovation is stimulated.

R&D Centers: Although most of the R&D centers are located in the MNCs' headquarters to avoid technology diffusion and keep their competitive advantage, MNCs are increasing their R&D expenditures overseas and establishing R&D centers in host countries (Braconier, Ekholm and Midelfart-Knarvik, 2001; UNCTAD, 2005). The capacity of generating knowledge in the host country is improved by participating in the R&D activities of MNCs.

Indirect Effects of FDI

Backward and Forward Linkages: A Backward linkage is the linkage between MNCs and suppliers, while a forward linkage occurs between

the MNCs and their customers and the companies that buy their products (Rodriguez-Clare, 1996). Backward linkages may help local suppliers promote their productivity by providing technical and information assistance (Belderbos, Capannelli and Fukao, 2001; Javorcik, 2004). In forward linkages local distributors and downstream firms can benefit from the MNC's knowledge to access higher-quality and/or lower-priced products.

Demonstration-imitation

effect: Due to technological differences between foreign and local firms, advanced technologies are introduced by foreign companies to the local industry. Local companies improve their productivity by watching and imitating the way foreign companies operate. Through learning by watching, local firms who are competitors of MNCs

improve their production processes through the disclosure of foreign advanced technology (Blomström, Kokko and Zejan, 1994).

Training effect: MNCs train their foreign partners, foreign buyers or suppliers, and local companies to maintain their competitiveness. Employees who are employed by foreign companies may diffuse knowledge, skills, and management practices learned to local companies through labor turnover or if they run their own businesses (Fosfuri and Saggi, 2002).

Competition effect: Because of the increased competition in the domestic market with the presence of MNCs, local firms are obligated to operate competently to avoid losing their market position (Bertschek 1995). Generally, this kind of spillover takes place at the intra-industry level. In other

words, companies in the same industry can be affected by competition imposed by MNCs with advanced technology.

Although FDI has the potential to improve productivity in the host country, the benefits are not guaranteed and are not independent of the conditions of each host country. The particular characteristics of the host country will determine the extent of those benefits. Specifically, an absorptive capability is required to cope with the new technology (Girma, 2003; Crespo and Fontoura, 2007). Sometimes, technologies MNCs bring to a host country are inappropriate for local companies and industries. Therefore, local companies are not able to improve their market position. In order to benefit from technology transfer, domestic firms and industries need to make certain investments. Spillover mainly depends on the absorptive capability of local

firms to become equal to the more developed foreign firms (Teece, 1977). When the technological gap between MNCs and local companies is significant, spillovers may not occur constructively.

At times, inward FDI can even worsen the host country's productivity. The technology transferred from the MNCs may have little influence on the host country's technological development and may even slow down the local productivity by restraining the local entrepreneurship since MNCs tend to dominate the local markets. There is also the possibility that the competition effect may have a negative impact on the local economy when local companies are not efficient enough to compete with foreign ones. Furthermore, local companies may become even less competitive and are eventually pushed out of business by foreign ones

(Cantwell 1995). Likewise, with FDI presence, the local productivity can decrease as the goal of those MNCs is to gain local market-share, by attracting demand from local competitors, which eventually decreases the local productivity (Aitken and Harrison, 1999). In addition, MNCs may tend to keep advanced technology and not transfer it to the host country in order to hold their monopoly status in technology (Ram and Zhang, 2002). Finally, foreign companies may draw the best workers from the local labor pool, leaving local companies with workers that are less skilled and less productive.

III. Model, Data and Methodology

We employ the widely adopted Cobb-Douglas production function model to test the relationship and the link between productivity and FDI.

Since changes in technology add value (Romer, 1990; Grossman and Helpman, 1991; Barro and Sala-i-Martin, 1995) to production, by incorporating technical factors associated with FDI and domestic factors into the original Cobb-Douglas production function, we incorporate the following form of the equation:

$$Y = f(L, K, H, R, F, S, G, E) \quad (1)$$

Where: Y (productivity) is taken as the current value-added in each sub-sectors of China's automotive industry.

L (input of labor) is measured by the total number of employees in each sub-sector.

K (Domestic capital stock) is defined by the current value of total domestic capital formation in each sub-sector. This suggested definition is in line with previous research, which

assumes that FDI leads to increases on the domestic stock of capital and production capacity (According to Egger and Pfaffermayr, 2001).

H (Human capital) is measured by the ratio of the number of technical staff to the annual average number of employees in each industry sub-sector. Human capital demonstrates the level of skill or education of employees.

R (Domestic technological efforts) is taken as the ratio of R&D expenditure by the total output in each sub-sector. Innovation stands for new ideas, methods and products that are introduced into production process or into the market, representing the technological capability of domestic economy.

F (Direct effects from FDI) is measured by the current value of FDI stock in each sub-sector. Since FDI transfers capital, technology and

management skills to their affiliates in host country, the greater value the foreign investment inflows will lead to the higher productivity.

S (Spillovers of FDI) is proxied by the ratio of output by foreign-invested enterprises in the sub-sectors of China's automotive industry to each sub-sector's total output.

G (Absorptive Capacity) is measured by the product of each sub-sector's human capital and FDI stock ($H * F$), which shows the ability of domestic firms to catch up with the technical knowledge of foreign firms and complementarities between domestic technological capacity and FDI.

E (Firm Size) is measured by the ratio of the total value of industrial output in each sub-sector to the number of firms in each sub-sector. Firm size stands for the economies of scale since it

is an important factor that affects the productivity in the automotive industry.

Based on the adopted production function, the following hypotheses are postulated:

H₁: The number of employees (L) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₂: value of domestic capital (K) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₃: the ratio of the number of technical staff to the annual average number of employees (H) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₄: the ratio of R&D expenditure to total output (R) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₅: the value of FDI stock (F) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₆: the ratio of output by foreign-invested enterprises to total output (S) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₇: the product of human capital and FDI (G) has a positive impact on each sub-sector's productivity in China's automotive industry.

H₈: the ratio of the value of industrial output to the number of firms (E) has a positive impact on each sub-sector's productivity in China's automotive industry.

It is expected that all of the individual independent variables has a positive impact on the productivity of the Chinese automotive industry. Hence,

a panel data set of each sub-sector in the industry is employed to test the model. The time period studied captures the period from 1999 to 2008. All the data were obtained from the Chinese Automotive Industry Yearbook 2000-2009, in which the industry is divided into five sub-sectors: auto-manufacturing, auto-assembling, motor-manufacturing, vehicle-engines, and vehicle-parts.

Consequently, a logarithmic model is employed to measure the elasticity of the impact of the independent variables on the dependent variable as described by the equation below:

$$\begin{aligned} \text{Ln}(Y_{it}) = & \alpha_i + \beta_1 \text{Ln}(L_{it}) + \\ & \beta_2 \text{Ln}(K_{it}) + \beta_3 \text{Ln}(H_{it}) + \beta_4 \text{Ln}(R_{it}) + \beta_5 \\ & \text{Ln}(F_{it}) + \beta_6 \text{Ln}(S_{it}) + \beta_7 \text{Ln}(G_{it}) + \\ & \beta_8 \text{Ln}(E_{it}) + \epsilon_{it} \quad (2) \end{aligned}$$

Where i and t denote the sub-sectors of the industry and time, respectively; α is the intercept and ϵ is the stochastic error term. The coefficients $\beta_1, \beta_2, \beta_3, \beta_4, \beta_5, \beta_6, \beta_7$ and β_8 show the percent change in $\text{Ln}(Y)$ related with percent change in variables L, K, H, R, F, S, G and E respectively.

Three statistical models are usually applied to estimate panel data sets: a pooled ordinary least squares model (POLS), a fixed effects model (FES), and a random effects model (RES). The main differences among these models are the assumptions, which are related to the intercepts and the error terms. Both the POLS model and the FES model are used to estimate equation (2). The RES model cannot be used in this research because the number of independent variables is larger than the number of cross-sections. Hence, the

Likelihood ratio (LR) test is used to determine which model is better (POLS or FES). We favor the FES estimation since the value of LR is significantly different from zero.

IV. Empirical Results

The empirical results from the POLS and FES model are summarized in the following (in Table 1??) table. As the table indicates, the FES model is preferred to the POLS model because of its large and significant LR-value. Therefore, the discussion is based only on the estimates of the FES model.

Results of Panel Data Estimations, 1999-2008 are as follows:

| Variable | POLS | FES |
|----------|------------------------|-------------------|
| Ln(L) | -0.7574(0.2250) | 0.0820(0.1976) |
| Ln(K) | 4.9277(0.2621) *** | -0.2162(0.2156) |
| Ln(H) | -1.1838(1.7500) | -2.0470(0.9068)** |
| Ln(R) | -3.7140(0.1249) *** | 0.2683(0.0820)*** |
| Ln(F) | -1.1926(1.7488) | -1.9853(0.8983)** |

| | | |
|-------------|-----------------|-------------------|
| Ln(S) | -0.3735(0.0681) | 0.0226(0.0691) |
| Ln(G) | 1.1930(1.7580) | 2.0187(0.9038)** |
| Ln(E) | -0.0844(0.0818) | 1.1076(0.1201)*** |
| Adjusted | | |
| R-squared | 0.9340 | 0.9813 |
| F-Statistic | 80.2458 | 214.9333 |
| Sample | | |
| Size (N) | 50 | 50 |
| Ln | | |
| likelihood | -8.874 | 27.2686 |

Notes: (1) Standard errors are in parentheses.

(2) *** significant at 1%, ** significant at 5%, * significant at 10%.

The results from the FES model display that domestic technological efforts Ln(R), absorptive capacity Ln(G) and firm size Ln(E) are positive as expected. Ln(R) and Ln(E) are statistically significant at a 1 % level and Ln(G) is statistically significant at a 5 % level. The coefficient for Ln(R) is positive and statistically significant at the 1 % level, indicating that R&D positively affects the productivity in China's automotive industry. The

magnitude of $\ln(R)$ may mean that when other variables are kept constant, a 1% increase in R&D increases productivity by 0.268 %.

The coefficient for $\ln(G)$ is positive and statistically significant at the 5 % level, showing that the absorptive capability positively affects productivity in China's automotive industry and that domestic human capital plays a role in capturing the benefits from FDI. In addition, The magnitude of $\ln(G)$ indicates that when other variables are kept constant, a 1% increase in absorptive capability will raise productivity by 2.018744 percent.

The magnitude of the coefficient $\ln(E)$ indicates that when other variables are kept constant, a 1% increase economy of scale will raise productivity by 1.108 %. The coefficient for $\ln(E)$ is positive and statistically

significant at the 1 % level, demonstrating that economy of scale positively affects productivity in China's automotive industry. This is an important finding and contribution to the emerging markets literature.

On the other hand and surprisingly, foreign direct investment $\ln(F)$ and human capital $\ln(H)$ are negative and statistically significant. Input of labor $\ln(L)$ and spillover in FDI $\ln(S)$ are positive as expected; however, they are statistically insignificant at different levels. Similarly, domestic capital stock $\ln(K)$ is negative but statistically insignificant at various levels as well.

Furthermore, the coefficient for $\ln(F)$ is negative and statistically significant at the 5 % level, demonstrating that direct FDI effects negatively affect productivity in China's

automotive industry. The magnitude of coefficient Ln (F) displays that when other variables are kept constant, a 1% increase in the direct FDI effect causes a decrease in productivity by 1.985 %. Hence, the result suggests that MNCs may not tend to transfer technology to host countries since they prefer to keep their monopoly status in technology (Ram and Zhang, 2002). This seems be the case in China, since the Chinese government only allows FDI in the form of Joint-Ventures in the automotive industry, thus MNCs may be discouraged to transfer their core technological capabilities.

The coefficient for Ln (H) is negative and statistically significant at the 5 % level, showing that human capital negatively affects productivity in China's automotive industry. The magnitude of Ln (H) shows that when

other variables are kept constant, 1% increase in human capital will decrease productivity by 2.047 percent. The result reflects the fact that compared to the total number of employees, the number of technically skilled employees is needed more in this industry since imported production lines are highly automated and only trained workers can operate them efficiently.

Although Ln (L), Ln (S) and Ln (K) are not significant at all levels, the coefficients of Ln (L) and Ln (S) are as expected indicating that these two factors contribute to productivity. However the coefficient of Ln (K) is negative, suggesting that the domestic capital negatively affects productivity in China's automotive industry. This result proposes that there may be capital market imperfection in this industry. This proposition may further be

supported by the status quo that SOEs have privileges to access capital and may be able to obtain subsidies from the government; thus compared to small and medium size enterprises (SMEs), they may lack the incentive to use capital efficiently.

Interestingly, our results contradict the FDI theories that suggest FDI has a positive impact on the host country's industrial productivity through both direct and indirect effects. This may be related to the competition effect and the unwillingness of core technology transfer. Based on the results, we can suggest that the Chinese government should not continue to place an ownership limit on FDI in the Chinese automotive industry.

Our results also indicate that the most influential factors to increase the productivity in the automotive industry

are the domestic technology effort, the domestic absorptive capability and the economy of scale. Hence the results suggest that it crucial for the Chinese government to continue to encourage R&D and consolidation to improve productivity level in the industry within the current development period. Furthermore, the results suggest that the domestic capital has a negative impact on the productivity in the industry indicating the existence of capital imperfection in the industry and suggesting that the government should treat SOEs and SMEs indifferently to improve their comparative advantages in order to compete not only in the domestic market, but also in international markets.

V. Conclusion

This paper focuses on the effects of FDI on the productivity of the

Chinese automotive industry by using a panel data set consisting of five sub-sectors over a period of ten years - from 1999 to 2008. Thus, the paper contributes to the empirical evidence concerning the FDI-productivity linkages in the economies of developing countries through a unique approach that emphasizes on a particular sector. In this paper, we model two channels, namely, the direct effects and spillovers through which FDI may affect local industries. We also test how human capital in the host country may behave together with FDI in influencing industrial productivity.

The results indicate an important finding and suggest that inward FDI plays a negative role in raising productivity in the automotive industry, which is one of the most crucial key sectors in Chinese economy. Yet,

productivity-augmenting effects from FDI on Chinese automotive industry do transpire neither through direct methods nor through spillovers. Hence, the results contradict the theory of FDI that MNCs play an important role to improve the host country's economy through introducing and transferring capital, advanced technologies and managerial skills. The results may also denote that governmental policies introduced to attract FDI are not effective enough to promote productivity.

Consequently, based on the results it is crucial to suggest that it may not be sensible for the Chinese government to keep imposing ownership limits on the inflow of FDI in the automotive industry as this practice decreases the productivity and does not allow the industry to benefit from direct effects of FDI. It is also recommended that the

Chinese government should treat SOEs and SMEs equally, stop giving privileges to SOEs, and encourage them to compete with the rest of the industry by incorporating efficiency and competency.

In conclusion, it is important to point out that due to data limitations, the time period studies in this paper is only 10 years. If the time span is extended to include the preceding years of 1990s and 1980s, the result would undoubtedly be very different. This is mostly attributable to the fact that the development of the Chinese automotive industry could have not been achieved without the participation of MNCs, especially in the early stages.

Finally, although in this study, it is shown that FDI has a negative impact on the productivity of the automotive industry, as a whole, it is likely that

some sub-sectors benefit from FDI and others do not. In order to clarify the benefits of FDI, and the ones that benefit from FDI, as well as to further understand the cause and effect relations in China, further study is required. It is, however, certain that the implications of our empirical results are valuable to government decision makers and joint-venture managers to promote their productivity and eventually enable them to compete globally.

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Appendix

Regression Output

| Dependent Variable: LN(?Y) | | | | |
|---|-------------|--------------------|-------------|--------|
| Method: Pooled Least Squares | | | | |
| Date: 03/19/11 Time: 21:49 | | | | |
| Sample: 1999 2008 | | | | |
| Included observations: 10 | | | | |
| Number of cross-sections used: 5 | | | | |
| Total panel (balanced) observations: 50 | | | | |
| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
| LN(?L) | 0.081986 | 0.197583 | 0.414946 | 0.6806 |
| LN(?K) | -0.216175 | 0.215563 | -1.002841 | 0.3225 |
| LN(?H) | -2.047024 | 0.906785 | -2.257452 | 0.0300 |
| LN(?R) | 0.268329 | 0.082033 | 3.270992 | 0.0023 |
| LN(?F) | -1.985265 | 0.898289 | -2.210052 | 0.0334 |
| LN(?S) | 0.022643 | 0.069100 | 0.327679 | 0.7450 |
| LN(?G) | 2.018744 | 0.903825 | 2.233555 | 0.0316 |
| LN(?E) | 1.107633 | 0.120074 | 9.224608 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 5.326694 | | | |
| _AUTOA--C | 6.834744 | | | |
| _MOTORM--C | 5.455930 | | | |
| _VE--C | 4.534370 | | | |
| _VP--C | 8.484960 | | | |
| R-squared | 0.985857 | Mean dependent var | 5.424238 | |
| Adjusted R-squared | 0.981271 | S.D. dependent var | 1.191332 | |
| S.E. of regression | 0.163040 | Sum squared resid | 0.983541 | |
| Ln likelihood | 27.26855 | F-statistic | 214.9333 | |
| Durbin-Watson stat | 2.168451 | Prob(F-statistic) | 0.000000 | |

Conduct, Interpret and Test the Regression

Estimation Command:

```

=====
EST(F,B,M=500,C=0.0001) LN(?Y)
LN(?L) LN(?K) LN(?H) LN(?R) LN(?F)
LN(?S) LN(?G) LN(?E)
    
```

Estimation Equations:

```

=====
LN(_AUTOMY) = C(9) +
C(1)*LN(_AUTOML) +
C(2)*LN(_AUTOMK) +
C(3)*LN(_AUTOMH) +
C(4)*LN(_AUTOMR) +
C(5)*LN(_AUTOMF) +
C(6)*LN(_AUTOMS) +
C(7)*LN(_AUTOMG) +
C(8)*LN(_AUTOME)

LN(_AUTOAY) = C(10) +
C(1)*LN(_AUTOAL) +
C(2)*LN(_AUTOAK) +
C(3)*LN(_AUTOAH) +
C(4)*LN(_AUTOAR) +
C(5)*LN(_AUTOAF) +
C(6)*LN(_AUTOAS) +
C(7)*LN(_AUTOAG) +
C(8)*LN(_AUTOAE)

LN(_MOTORMY) = C(11) +
C(1)*LN(_MOTORML) +
C(2)*LN(_MOTORMK) +
C(3)*LN(_MOTORMH) +
C(4)*LN(_MOTORMR) +
C(5)*LN(_MOTORMF) +
C(6)*LN(_MOTORMS) +
C(7)*LN(_MOTORMG) +
C(8)*LN(_MOTORME)
    
```

$$\begin{aligned} \text{LN_VEY} = & C(12) + C(1)*\text{LN_VEL} \\ & + C(2)*\text{LN_VEK} + C(3)*\text{LN_VEH} + \\ & C(4)*\text{LN_VER} + C(5)*\text{LN_VEF} + \\ & C(6)*\text{LN_VES} + C(7)*\text{LN_VEG} + \\ & C(8)*\text{LN_VEE} \end{aligned}$$

$$\begin{aligned} \text{LN_VPY} = & C(13) + C(1)*\text{LN_VPL} + \\ & C(2)*\text{LN_VPK} + C(3)*\text{LN_VPH} + \\ & C(4)*\text{LN_VPR} + C(5)*\text{LN_VPF} + \\ & C(6)*\text{LN_VPS} + C(7)*\text{LN_VPG} + \\ & C(8)*\text{LN_VPE} \end{aligned}$$

Substituted Coefficients:

=====

$$\begin{aligned} \text{LN_AUTOMY} = & 5.326693765 + \\ & 0.08198634807*\text{LN_AUTOML} - \\ & 0.216175275*\text{LN_AUTOMK} - \\ & 2.047023508*\text{LN_AUTOMH} + \\ & 0.268329243*\text{LN_AUTOMR} - \\ & 1.985264718*\text{LN_AUTOMF} + \\ & 0.02264279557*\text{LN_AUTOMS} + \\ & 2.018744196*\text{LN_AUTOMG} + \\ & 1.107632601*\text{LN_AUTOME} \end{aligned}$$

$$\begin{aligned} \text{LN_AUTOAY} = & 6.834743645 + \\ & 0.08198634807*\text{LN_AUTOAL} - \\ & 0.216175275*\text{LN_AUTOAK} - \\ & 2.047023508*\text{LN_AUTOAH} + \\ & 0.268329243*\text{LN_AUTOAR} - \\ & 1.985264718*\text{LN_AUTOAF} + \\ & 0.02264279557*\text{LN_AUTOAS} + \\ & 2.018744196*\text{LN_AUTOAG} + \\ & 1.107632601*\text{LN_AUTOAE} \end{aligned}$$

$$\begin{aligned} \text{LN_MOTORMY} = & 5.455930021 + \\ & 0.08198634807*\text{LN_MOTORML} - \\ & 0.216175275*\text{LN_MOTORMK} - \\ & 2.047023508*\text{LN_MOTORMH} + \\ & 0.268329243*\text{LN_MOTORMR} - \\ & 1.985264718*\text{LN_MOTORMF} + \\ & 0.02264279557*\text{LN_MOTORMS} + \\ & 2.018744196*\text{LN_MOTORMG} + \\ & 1.107632601*\text{LN_MOTORME} \end{aligned}$$

$$\begin{aligned} \text{LN_VEY} = & 4.534369629 + \\ & 0.08198634807*\text{LN_VEL} - \\ & 0.216175275*\text{LN_VEK} - \\ & 2.047023508*\text{LN_VEH} + \\ & 0.268329243*\text{LN_VER} - \\ & 1.985264718*\text{LN_VEF} + \\ & 0.02264279557*\text{LN_VES} + \\ & 2.018744196*\text{LN_VEG} + \\ & 1.107632601*\text{LN_VEE} \end{aligned}$$

$$\begin{aligned} \text{LN_VPY} = & 8.484959842 + \\ & 0.08198634807*\text{LN_VPL} - \\ & 0.216175275*\text{LN_VPK} - \\ & 2.047023508*\text{LN_VPH} + \\ & 0.268329243*\text{LN_VPR} - \\ & 1.985264718*\text{LN_VPF} + \\ & 0.02264279557*\text{LN_VPS} + \\ & 2.018744196*\text{LN_VPG} + \\ & 1.107632601*\text{LN_VPE} \end{aligned}$$

$$\begin{aligned} \text{Ln}(\bar{Y}_{it}) = & \alpha_i + 0.082*\text{Ln}(L_{it}) - 0.2162*\text{Ln}(K_{it}) - \\ & 2.047*\text{Ln}(H_{it}) + 0.2683*\text{Ln}(R_{it}) - 1.9853*\text{Ln}(F_{it}) + \\ & 0.0264*\text{Ln}(S_{it}) + 2.0187*\text{Ln}(G_{it}) + 1.1076*\text{Ln}(E_{it}) \end{aligned}$$

| | (0.1976) | (0.2156) | (0.9068) | (0.082) | (0.8983) | (0.0691) | (0.9038) | (0.1201) |
|---|----------|----------|----------|---------|----------|----------|----------|----------|
| T | 0.415 | -1.0028 | -2.2575 | 3.271 | -2.2101 | 0.3277 | 2.2336 | 9.2246 |

| Fixed Effects | α_i |
|-----------------------------------|------------|
| _AUTOM--C | 5.326694 |
| _AUTOA--C | 6.834744 |
| _MOTORM--C | 5.455930 |
| _VE--C | 4.534370 |
| _VP--C | 8.484960 |
| N=50 $\bar{R}^2=0.9813$ DW=2.1685 | |

Descriptive Statistics

| | ?Y | ?L | ?K | ?H | ?R | ?F | ?S | ?G | ?E |
|---------|-------|------|-------|------|------|------|------|------|------|
| Mean | 446.9 | 35.2 | 1767. | 0.12 | 0.01 | 28.7 | 0.23 | 3.03 | 11.4 |
| | 340 | 4000 | 664 | 0606 | 8306 | 6000 | 7105 | 5949 | 4342 |
| Sum | 2234 | 1762 | 8838 | 6.03 | 0.91 | 1438 | 11.8 | 151. | 572. |
| | 6.70 | .000 | 3.20 | 0287 | 5314 | .000 | 5526 | 7974 | 1709 |
| Median | 196.2 | 20.6 | 702.1 | 0.11 | 0.01 | 8.20 | 0.24 | 0.76 | 3.64 |
| | 000 | 5000 | 000 | 1492 | 5852 | 0000 | 8120 | 1331 | 0682 |
| Maxim | 2135. | 101. | 7549. | 0.47 | 0.06 | 167. | 0.48 | 20.3 | 88.6 |
| um | 300 | 9000 | 000 | 4970 | 8710 | 8000 | 4372 | 8165 | 7830 |
| Minim | 19.80 | 5.00 | 224.8 | 0.00 | 0.00 | 0.10 | 0.00 | 0.01 | 0.39 |
| um | 000 | 0000 | 000 | 8881 | 4190 | 0000 | 3071 | 8220 | 5206 |
| Sum | 1497 | 3975 | 1.62 | 0.18 | 0.00 | 7866 | 1.06 | 925. | 1747 |
| Sq. | 1427 | 4.08 | E+08 | 1494 | 4428 | 6.42 | 1033 | 2253 | 3.68 |
| Dev. | | | | | | | | | |
| Std. | 552.7 | 28.4 | 1819. | 0.06 | 0.00 | 40.0 | 0.14 | 4.34 | 18.8 |
| Dev. | 561 | 8346 | 152 | 0860 | 9506 | 6791 | 7152 | 5359 | 8401 |
| Skewne | 1.656 | 0.73 | 1.371 | 4.24 | 3.19 | 1.66 | 0.04 | 1.95 | 2.37 |
| ss | 683 | 9366 | 115 | 4056 | 5501 | 4027 | 0006 | 2353 | 7623 |
| Kurtosi | 4.715 | 2.32 | 3.891 | 24.9 | 17.0 | 5.14 | 1.67 | 6.92 | 8.26 |
| s | 386 | 9692 | 937 | 9973 | 6770 | 5007 | 5742 | 1447 | 2093 |
| Jarque- | 29.00 | 5.49 | 17.32 | 1158 | 497. | 32.6 | 3.66 | 63.8 | 104. |
| Bera | 197 | 1579 | 369 | .409 | 3856 | 6043 | 6792 | 0100 | 7958 |
| Probabi | 0.000 | 0.06 | 0.000 | 0.00 | 0.00 | 0.00 | 0.15 | 0.00 | 0.00 |
| lity | 001 | 4198 | 173 | 0000 | 0000 | 0000 | 9870 | 0000 | 0000 |
| Observ | | | | | | | | | |
| ations | 50 | 50 | 50 | 50 | 50 | 50 | 50 | 50 | 50 |
| Cross | | | | | | | | | |
| section | | | | | | | | | |
| s | 5 | 5 | 5 | 5 | 5 | 5 | 5 | 5 | 5 |

F-Testing

Hypothesis:
 $H_0: \beta_1 = \beta_2 = \beta_3 = \beta_4 = \beta_5 = \beta_6 = \beta_7 = \beta_8 = 0$
 $H_a: \text{at least one } \beta_t \neq 0$

In the regression output, the probability of F-statistic=0, so we can reject H_0 at 1% level. Therefore, the overall fit of the equation is statistically significant at 1% level.

Hypothesis Testing

1. Test the sign and significance of Ln(L) at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{Ln(L)} \leq 0, H_a: \beta_{Ln(L)} > 0$

The slop coefficient of Ln(L) is positive as we expected. The P-value is 0.3403 for one tail, which is insignificant at 1%, 5% and 10% level. Therefore, we cannot reject H_0 at all levels.

2. Test the sign and significance of Ln(K) at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{Ln(K)} \leq 0, H_a: \beta_{Ln(K)} > 0$

The slope coefficient of Ln(K) is negative as we unexpected. The P-value is 0.1613 for one tail, which is insignificant at 1%, 5% and 10% level. Thus, we cannot reject H_0 at all levels.

3. Test the sign and significance of Ln(H) at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{Ln(H)} \leq 0, H_a: \beta_{Ln(H)} > 0$

The slope coefficient of Ln(H) is negative as we unexpected. The P-value is 0.015 for one tail, which is insignificant at 1% level of confidence, however is significant at 5% and 10% level of

confidence. Therefore, we cannot reject H_0 at all levels.

4. Test the sign and significance of $\ln(R)$ at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{\ln(R)} \leq 0, H_a: \beta_{\ln(R)} > 0$

The slope coefficient of $\ln(R)$ is positive as we expected. The P-value is 0.00125 for one tail, which is significant at 1%, 5% and 10% level. As a result, we can reject H_0 at all levels.

5. Test the sign and significance of $\ln(F)$ at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{\ln(F)} \leq 0, H_a: \beta_{\ln(F)} > 0$

The slope coefficient of $\ln(F)$ is negative as we unexpected. The P-value is 0.0167 for one tail, which is insignificant at 1% confident level but 5% and 10% level. Therefore, we cannot reject H_0 at all levels.

6. Test the sign and significance of $\ln(S)$ at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{\ln(S)} \leq 0, H_a: \beta_{\ln(S)} > 0$

The slope coefficient of $\ln(S)$ is positive as we expected. The P-value is 0.3725 for one tail, which is insignificant at 1%, 5% and 10% level. As a result, we cannot reject H_0 at all levels.

7. Test the sign and significance of $\ln(G)$ at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{\ln(G)} \leq 0, H_a: \beta_{\ln(G)} > 0$

The slope coefficient of $\ln(G)$ is positive we expected. The P- value is 0.0158 for one tail, which is insignificant at 1% level but significant at 5% and 10% level. Therefore, we cannot reject H_0 at 1% level but we can reject H_0 at 5% and 10% level.

8. Test the sign and significance of $\ln(E)$ at the 1%, 5% and 10% level.

Hypothesis: $H_0: \beta_{\ln(E)} \leq 0, H_a: \beta_{\ln(E)} > 0$

The slope coefficient of $\ln(E)$ is positive as we expected. The P-value is 0 for one tail, which is significant at 1%, 5% and 10% level. Thus, we can reject H_0 at all levels.

Irrelevant Variables and Omitted Variables Testing $\ln(L)$

Dependent Variable: $\ln(?Y)$

Method: Pooled Least Squares

Date: 03/20/11 Time: 00:00

Sample: 1999 2008

Included observations: 10

Number of cross-sections used: 5

Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|---------------|-------------|------------|-------------|--------|
| $\ln(?K)$ | -0.171756 | 0.185055 | -0.928135 | 0.3592 |
| $\ln(?H)$ | -2.077665 | 0.893875 | -2.324336 | 0.0256 |
| $\ln(?R)$ | 0.273355 | 0.080245 | 3.406479 | 0.0016 |
| $\ln(?F)$ | -1.996759 | 0.888028 | -2.248531 | 0.0304 |
| $\ln(?S)$ | 0.029097 | 0.066589 | 0.436968 | 0.6646 |
| $\ln(?G)$ | 2.028649 | 0.893615 | 2.270161 | 0.0290 |
| $\ln(?E)$ | 1.089716 | 0.110815 | 9.833683 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 5.337568 | | | |
| _AUTOA--C | 6.795396 | | | |

| | | | |
|--------------------|----------|--------------------|----------|
| _MOTORM--C | 5.406406 | | |
| _VE--C | 4.440478 | | |
| _VP--C | 8.481541 | | |
| R-squared | 0.985792 | Mean dependent var | 5.424238 |
| Adjusted R-squared | 0.981679 | S.D. dependent var | 1.191332 |
| S.E. of regression | 0.161255 | Sum squared resid | 0.988118 |
| Ln likelihood | 27.15249 | F-statistic | 239.6784 |
| Durbin-Watson stat | 2.208052 | Prob(F-statistic) | 0.000000 |

1. Theory: as hypothesis 1 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(L) for one tail is 0.3403, which is insignificant at all levels. Thus, it should be an irrelevant variable.
3. Adjusted R-squared: the \bar{R}^2 increased slightly from 0.9813 to 0.9817. It indicates that Ln(L) should not belong to this equation.
4. Bias: with Ln(L) removed, all coefficients changed slightly. Therefore, it should be an irrelevant variable.

To sum up, the variable Ln(L) should belong to this equation.

Testing Ln(K)

Dependent Variable: LN(?Y)
 Method: Pooled Least Squares
 Date: 03/20/11 Time: 00:06
 Sample: 1999 2008
 Included observations: 10

Number of cross-sections used: 5
 Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|----------|-------------|------------|-------------|--------|
| LN(?L) | -0.016412 | 0.171511 | -0.095691 | 0.9243 |
| LN(?H) | -2.141613 | 0.901934 | -2.374468 | 0.0227 |
| LN(?R) | 0.275813 | 0.081699 | 3.375962 | 0.0017 |
| LN(?F) | -2.031471 | 0.897173 | -2.264301 | 0.0293 |
| LN(?S) | 0.049296 | 0.063789 | 0.772801 | 0.4444 |
| LN(?G) | 2.055809 | 0.903137 | 2.276298 | 0.0285 |
| LN(?E) | 0.999257 | 0.052339 | 19.09212 | 0.0000 |

Fixed Effects

| | | | |
|--------------------|----------|--------------------|----------|
| _AUTOM--C | 4.274776 | | |
| _AUTOA--C | 5.751558 | | |
| _MOTORM--C | 4.446693 | | |
| _VE--C | 3.567653 | | |
| _VP--C | 7.204965 | | |
| R-squared | 0.985473 | Mean dependent var | 5.424238 |
| Adjusted R-squared | 0.981268 | S.D. dependent var | 1.191332 |
| S.E. of regression | 0.163053 | Sum squared resid | 1.010274 |
| Ln likelihood | 26.59810 | F-statistic | 234.3462 |
| Durbin-Watson stat | 2.222591 | Prob(F-statistic) | 0.000000 |

1. Theory: as hypothesis 2 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(K) for one tail is 0.1613, which is significant at 5% and 10% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9812. It indicates that Ln(K) should be a relevant variable.

4. Bias: with Ln(K) removed, some coefficients changed significantly. Therefore, it should belong to the equation.

To sum up, the variable Ln(K) should belong to this equation.

Testing Ln(H)

Dependent Variable: LN(?Y)
 Method: Pooled Least Squares
 Date: 03/20/11 Time: 00:07
 Sample: 1999 2008
 Included observations: 10
 Number of cross-sections used: 5
 Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--------------------|-------------|--------------------|-------------|--------|
| LN(?L) | 0.118310 | 0.207269 | 0.570803 | 0.5715 |
| LN(?K) | -0.266792 | 0.225653 | -1.182314 | 0.2444 |
| LN(?R) | 0.262277 | 0.086295 | 3.039304 | 0.0043 |
| LN(?F) | 0.022470 | 0.132820 | 0.169173 | 0.8666 |
| LN(?S) | 0.005416 | 0.072285 | 0.074930 | 0.9407 |
| LN(?G) | 0.001194 | 0.141792 | 0.008418 | 0.9933 |
| LN(?E) | 1.132610 | 0.125842 | 9.000260 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 5.577782 | | | |
| _AUTOA--C | 7.053500 | | | |
| _MOTORM--C | 5.671369 | | | |
| _VE--C | 4.766150 | | | |
| _VP--C | 8.775257 | | | |
| R-squared | 0.983909 | Mean dependent var | 5.424238 | |
| Adjusted R-squared | 0.979252 | S.D. dependent var | 1.191332 | |
| S.E. of regression | 0.171603 | Sum squared resid | 1.119006 | |
| Ln likelihood | 24.04263 | F-statistic | 211.2395 | |
| Durbin-Watson stat | 2.113126 | Prob(F-statistic) | 0.000000 | |

1. Theory: as hypothesis 3 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(H) for one tail is 0.015, which is significant at 5% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9792. It indicates that Ln(H) should be a relevant variable.
4. Bias: with Ln(H) removed, most of the coefficients changed significantly. Therefore, it should belong to the equation.

To sum up, the variable Ln(H) should belong to this equation.

Testing Ln(R)

Dependent Variable: LN(?Y)
 Method: Pooled Least Squares
 Date: 03/20/11 Time: 00:08
 Sample: 1999 2008
 Included observations: 10
 Number of cross-sections used: 5
 Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|---------------|-------------|------------|-------------|--------|
| LN(?L) | 0.177401 | 0.218942 | 0.810263 | 0.4228 |
| LN(?K) | -0.280316 | 0.240510 | -1.165505 | 0.2511 |
| LN(?H) | -1.950082 | 1.015399 | -1.920507 | 0.0623 |
| LN(?F) | -1.673329 | 1.000736 | -1.672099 | 0.1027 |
| LN(?S) | 0.138016 | 0.066573 | 2.073150 | 0.0450 |
| LN(?G) | 1.694066 | 1.006501 | 1.683124 | 0.1006 |
| LN(?E) | 1.176640 | 0.132435 | 8.884653 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 3.814688 | | | |
| _AUTOA--C | 5.765333 | | | |
| _MOTORM--C | 4.095647 | | | |

| | | | |
|--------------------|----------|--------------------|----------|
| -C | | | |
| _VE--C | 3.213681 | | |
| _VP--C | 7.147873 | | |
| R-squared | 0.981768 | Mean dependent var | 5.424238 |
| Adjusted R-squared | 0.976490 | S.D. dependent var | 1.191332 |
| S.E. of regression | 0.182667 | Sum squared resid | 1.267954 |
| Ln likelihood | 20.91854 | F-statistic | 186.0192 |
| Durbin-Watson stat | 2.080478 | Prob(F-statistic) | 0.000000 |

1. Theory: as hypothesis 4 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(R) for one tail is 0.00125, which is significant at 1% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9765. It indicates that Ln(R) should be a relevant variable.
4. Bias: with Ln(R) removed, some coefficients changed significantly. Therefore, it should belong to the equation.

To sum up, the variable Ln(R) should belong to this equation.

Testing Ln(F)

Dependent Variable: LN(?Y)
 Method: Pooled Least Squares
 Date: 03/20/11 Time: 00:09
 Sample: 1999 2008
 Included observations: 10
 Number of cross-sections used: 5

Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|---------------|-------------|------------|-------------|--------|
| LN(?L) | 0.095451 | 0.207337 | 0.460368 | 0.6479 |
| LN(?K) | -0.240611 | 0.226014 | -1.064585 | 0.2938 |
| LN(?H) | -0.062855 | 0.133738 | -0.469985 | 0.6411 |
| LN(?R) | 0.249082 | 0.085637 | 2.908584 | 0.0060 |
| LN(?S) | 0.024155 | 0.072543 | 0.332982 | 0.7410 |
| LN(?G) | 0.022112 | 0.027982 | 0.790205 | 0.4343 |
| LN(?E) | 1.123657 | 0.125831 | 8.929871 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 5.363287 | | | |
| _AUTOA--C | 6.871578 | | | |
| _MOTORM--C | 5.468140 | | | |
| _VE--C | 4.552025 | | | |
| _VP--C | 8.549084 | | | |

| | | | |
|--------------------|----------|--------------------|----------|
| R-squared | 0.983990 | Mean dependent var | 5.424238 |
| Adjusted R-squared | 0.979356 | S.D. dependent var | 1.191332 |
| S.E. of regression | 0.171171 | Sum squared resid | 1.113377 |
| Ln likelihood | 24.16871 | F-statistic | 212.3250 |
| Durbin-Watson stat | 2.106214 | Prob(F-statistic) | 0.000000 |

1. Theory: as hypothesis 5 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(F) for one tail is 0.0167, which is significant at 5% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9794. It indicates that Ln(F) should be a relevant variable.
4. Bias: with Ln(F) removed, some coefficients changed significantly.

Therefore, it should belong to the equation.

To sum up, the variable Ln(F) should belong to this equation.

Testing Ln(S)

Dependent Variable: LN(?Y)

Method: Pooled Least Squares

Date: 03/20/11 Time: 00:10

Sample: 1999 2008

Included observations: 10

Number of cross-sections used: 5

Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--------------------|-------------|--------------------|-------------|--------|
| LN(?L) | 0.096561 | 0.190237 | 0.507583 | 0.6147 |
| LN(?K) | -0.243344 | 0.196629 | -1.237580 | 0.2235 |
| LN(?H) | -2.014210 | 0.890591 | -2.261656 | 0.0295 |
| LN(?R) | 0.282050 | 0.069708 | 4.046180 | 0.0002 |
| LN(?F) | -1.988180 | 0.887632 | -2.239870 | 0.0310 |
| LN(?G) | 2.025247 | 0.892932 | 2.268088 | 0.0291 |
| LN(?E) | 1.115672 | 0.116151 | 9.605330 | 0.0000 |
| Fixed Effects | | | | |
| _AUTOM--C | 5.568539 | | | |
| _AUTOA--C | 7.031517 | | | |
| _MOTORM--C | 5.678094 | | | |
| _VE--C | 4.760362 | | | |
| _VP--C | 8.737241 | | | |
| R-squared | 0.985816 | Mean dependent var | 5.424238 | |
| Adjusted R-squared | 0.981711 | S.D. dependent var | 1.191332 | |
| S.E. of regression | 0.161114 | Sum squared resid | 0.986395 | |
| Ln likelihood | 27.19611 | F-statistic | 240.1030 | |
| Durbin-Watson stat | 2.178565 | Prob(F-statistic) | 0.000000 | |

1. Theory: as hypothesis 6 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(S) for one tail is 0.3725, which is significant at 5% and 10% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 increased slightly from 0.9813 to 0.9817. It indicates that Ln(S) should be an irrelevant variable.
4. Bias: with Ln(S) removed, some of the coefficients changed significantly. Therefore, it should belong to the equation.

To sum up, the variable Ln(S) should belong to this equation.

Testing Ln(G)

Dependent Variable: LN(?Y)

Method: Pooled Least Squares

Date: 03/20/11 Time: 00:10

Sample: 1999 2008

Included observations: 10

Number of cross-sections used: 5

Total panel (balanced) observations: 50

| Variable | Coefficien t | Std. Error | t-Statistic | Prob. |
|----------|--------------|------------|-------------|--------|
| LN(?L) | 0.093641 | 0.207622 | 0.451018 | 0.6545 |
| LN(?K) | -0.235864 | 0.226405 | -1.041782 | 0.3041 |
| LN(?H) | -0.044293 | 0.142075 | -0.311762 | 0.7569 |
| LN(?R) | 0.248207 | 0.085709 | 2.895915 | 0.0062 |

| | | | | |
|--------------------|----------|--------------------|---------|--------|
| LN(?F) | 0.020240 | 0.02784 | 0.72687 | 0.4718 |
| | | 6 | 8 | |
| LN(?S) | 0.026032 | 0.07261 | 0.35846 | 0.7220 |
| | | 9 | 8 | |
| LN(?E) | 1.123061 | 0.12600 | 8.91251 | 0.0000 |
| | | 9 | 7 | |
| Fixed Effects | | | | |
| _AUTOM--C | 5.331692 | | | |
| _AUTOA--C | 6.842342 | | | |
| _MOTORM--C | 5.436364 | | | |
| _VE--C | 4.522267 | | | |
| _VP--C | 8.517270 | | | |
| R-squared | 0.983950 | Mean dependent var | 5.42423 | |
| | | | 8 | |
| Adjusted R-squared | 0.979305 | S.D. dependent var | 1.19133 | |
| | | | 2 | |
| S.E. of regression | 0.171384 | Sum squared resid | 1.11615 | |
| | | | 3 | |
| Ln likelihood | 24.10645 | F-statistic | 211.788 | |
| | | | 3 | |
| Durbin-Watson stat | 2.112292 | Prob(F-statistic) | 0.00000 | |
| | | | 0 | |

1. Theory: as hypothesis 7 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(G) for one tail is 0.0158, which is significant at 5% level. Thus, it should belong to the equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9793. It indicates that Ln(G) should be a relevant variable.
4. Bias: with Ln(G) removed, some coefficients changed significantly. Therefore, it should belong to the equation.

To sum up, the variable Ln(G) should belong to this equation.

Testing Ln(E)

Dependent Variable: LN(?Y)
 Method: Pooled Least Squares
 Date: 03/20/11 Time: 00:11
 Sample: 1999 2008
 Included observations: 10
 Number of cross-sections used: 5
 Total panel (balanced) observations: 50

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--------------------|-------------|--------------------|-------------|--------|
| LN(?L) | -0.573421 | 0.330473 | -1.735151 | 0.0908 |
| LN(?K) | 1.573491 | 0.168411 | 9.343153 | 0.0000 |
| LN(?H) | -2.817801 | 1.618479 | -1.741018 | 0.0898 |
| LN(?R) | 0.401285 | 0.144755 | 2.772167 | 0.0086 |
| LN(?F) | -2.485638 | 1.607226 | -1.546539 | 0.1303 |
| LN(?S) | 0.152894 | 0.121248 | 1.261006 | 0.2150 |
| LN(?G) | 2.498364 | 1.617406 | 1.544673 | 0.1307 |
| Fixed Effects | | | | |
| _AUTOM--C | -2.902205 | | | |
| _AUTOA--C | -2.212636 | | | |
| _MOTORM--C | -2.497163 | | | |
| _VE--C | -2.935549 | | | |
| _VP--C | -2.378402 | | | |
| R-squared | 0.953332 | Mean dependent var | 5.424238 | |
| | | | | |
| Adjusted R-squared | 0.939823 | S.D. dependent var | 1.191332 | |
| | | | | |
| S.E. of regression | 0.292247 | Sum squared resid | 3.245509 | |
| | | | | |
| Ln likelihood | -2.578155 | F-statistic | 70.56896 | |
| | | | | |
| Durbin-Watson stat | 1.441169 | Prob(F-statistic) | 0.000000 | |
| | | | | |

1. Theory: as hypothesis 8 mentioned, this variable is sound theoretically.
2. T-test: The P-value of Ln(E) for one tail is 0, which is significant at all levels. Thus, it should belong to this equation.
3. Adjusted R-squared: the \bar{R}^2 decreased slightly from 0.9813 to 0.9398. It indicates that Ln(E) should be relevant variable.
4. Bias: with Ln(E) removed, all coefficients changed significantly. Thus, it should belong to the equation.

To sum up, the variable Ln(E) should belong to this equation.

Serial correlation

Durbin-Watson testing

The D-value from the regression output is 2.1685, N=50, and K=8.

There is potential of serial-correlation, since the data set contains time-series data.

$H_0: \rho=0$ (no serial correlation), $H_a: \rho \neq 0$ (serial correlation)

$$d_L = 1.2, d_U = 1.93$$

Since $4 - d_L = 2.8 > D\text{-value} = 2.1685 > 4 - d_U = 2.07$, the result is inconclusive, we cannot be sure if there exists serial-correlation in the equation at 5% level. Thus, General Least Square model is not required.