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Title: Group Adaptation, Formal Darwinism and Contextual Analysis

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Running Title: Group Adaptation, Formal Darwinism and Contextual Analysis

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1	Abstract
2	We consider the question: under what circumstances can the concept
3	of adaptation be applied to groups, rather than individuals? Gardner
4	and Grafen (2009) develop a novel approach to this question, building on
5	Grafen's 'formal Darwinism' project, which defines adaptation in terms
6	of links between evolutionary dynamics and optimization. They conclude
7	that only clonal groups, and to a lesser extent groups in which reproduc-
8	tive competition is repressed, can be considered as adaptive units. We
9	re-examine the conditions under which the selection-optimization links
10	hold at the group level. We focus on an important distinction between
11	two ways of understanding the links, which have different implications
12	regarding group adaptationism. We show how the formal Darwinism ap-
13	proach can be reconciled with G.C. Williams' famous analysis of group
14	adaptation, and we consider the relationships between group adaptation,
15	the Price equation approach to multi-level selection, and the alternative
16	approach based on contextual analysis.

Keywords: adaptation, group adaptation, superorganism, Price's equation, optimality, contextual analysis, G. C. Williams, formal Darwinism, group
selection

Group Adaptation, Formal Darwinism and Contextual Analysis

²³ 1 Introduction

Evolutionary biologists usually apply the concept of adaptation to individual 24 organisms. However it has long been recognised that in principle, groups might 25 also exhibit adaptations. The idea of group adaptation, and the associated con-26 cept of a 'superorganism', were famously criticised by G.C. Williams (1966), but 27 have since been revived by proponents of 'multi-level selection' (Sober & Wilson 28 1998; Seeley 1989, 1997; Hölldobler & Wilson 2009). Progress on this topic has 29 been hampered by unclarity about how exactly 'group adaptation' should be 30 defined, how it relates to 'group selection', and the conditions under which it 31 can evolve. Gardner and Grafen (2009) (hereafter G&G) make a remarkable 32 contribution by bringing mathematical precision to these issues, with striking 33 results. They do this by applying Grafen's 'formal Darwinism' project, (Grafen 34 2002, 2006, 2008), which provides a general framework for understanding the 35 concept of adaptation, to groups. 36

Our aim here is to take further the analysis of group adaptation, using a similar methodology to G&G. We recognise the merits of making the concept of group adaptation precise, and share their view that the formal Darwinism project offers the best way to do this. However, G&G's analysis leaves open a number of issues. In particular, it is unclear how the concept of group adaptation they articulate relates to G.C. Williams' (1966) well-known analysis of the concept (cf. Sober and Wilson 2011).

Our discussion falls into three parts. Firstly, we explore a subtle difference between two ways of defining adaptation using the formal Darwinism machinery, one used by G&G, the other by Grafen in his earlier papers. The two definitions have different implications in general; and as applied to groups, they differ on whether clonality, or repression of within-group competition, represent the clearest cases of group adaptation.

Secondly, we study how the formal Darwinism approach can be reconciled with G.C. Williams' distinction between 'group adaptation' and 'fortuitous group benefit'. The former refers to a group feature that evolved *because* it benefits the group, the latter to a group feature that happens to benefit the group but did not evolve for that reason. (Thus Williams famously contrasted a 'herd of fleet deer' with a 'fleet herd of deer'.) Many biologists regard this dis-

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tinction as crucial, so it is of some interest to see whether the formal Darwinism
 approach to group adaptation can accommodate it.

Thirdly and relatedly, we consider the relation between the 'Price equation' 58 approach to multi-level selection and the alternative approach based on 'con-59 textual analysis'. These approaches constitute alternative ways of partitioning 60 the total evolutionary change in a structured population into components cor-61 responding to distinct levels of selection. G&G say that their analysis 'has 62 identified Price's between-group selection as the driver of group adaptation', 63 and thus favour the Price approach (p. 667). We show that the contextual 64 approach can also supply a formal definition of group adaptation. 65

⁶⁶ 2 The 'maximizing agent' analogy

G&G's analysis of group adaptation draws on Grafen's 'Formal Darwinism' 67 project, which aims to connect optimization and natural selection in a precise 68 way, thus formally justifying the intuitive idea that selection leads to organismic 69 design (Grafen 2002, 2006, 2008). Grafen's approach is to use a fully explicit 70 definition of optimization, then to prove links between optimality and evolution-71 ary dynamics. The notion of optimization is captured by an 'objective function' 72 that maps an agent's phenotype to its 'fitness' (for some measure of fitness); 73 if an agent achieves the maximum value of this function, they are said to 'be-74 have optimally'. The links state logical connections between the optimality or 75 otherwise of agents' behaviour and the operation of natural selection. 76

In Grafen's original papers, the 'agents' are taken to be individual organ-77 isms; this is natural because individuals are usually treated as the bearers of 78 adaptations in biology. With this interpretation, the links capture the sense in 79 which natural selection leads individuals to be adaptive units, just as Darwin 80 originally argued. G&G investigate what happens when the 'agents' in Grafen's 81 analysis are instead taken to be groups; their aim is to see whether, and in 82 what circumstances, whole groups can legitimately be considered as adaptive 83 units, or 'maximizing agents'. Their main conclusion is that these circumstances 84 are relatively rare, because the required links between optimality and natural 85 selection only hold under fairly stringent conditions. 86

To understand G&G's argument, the optimality / selection links must be 87 laid out explicitly (see Table 1). The first link says that if all agents behave 88 optimally, there is no 'scope' for selection, i.e. no gene will change in (expected) 89 frequency. This makes good sense: if all agents achieve maximum fitness, the 90 fitness variance in the population is zero, so no selection will occur. The second 91 link says that if all agents behave optimally, there is no 'potential' for positive 92 selection, which means that no introduced mutant will spread. This also makes 93 sense: if all agents achieve maximum fitness, then no mutant can do better. The 94 third link says that if all agents behave suboptimally, but equally so, there is no 95 scope for selection. Again this makes sense, given that selection requires variance 96 in fitness. The fourth link says that if all agents behave suboptimally, but 97 equally so, then there is potential for positive selection. This is also intuitive, 98

1. If all agents behave optimally, there is no scope for selection.

If all agents behave optimally, there is no potential for positive selection.
 If all agents behave suboptimally, but equally so, there is no scope for selection.

4. If all agents behave suboptimally, but equally so, then there is potential for positive selection.

5. If agents vary in their optimality, then there is scope for selection, and the change in the frequency of any gene is given by the covariance between the frequency of that gene in an agent and the agent's relative fitness.

Table 1: The selection / optimality links

⁹⁹ since a mutant phenotype which achieves a higher fitness than the incumbents¹⁰⁰ will spread in the population.

The fifth link is slightly different, in that it describes what will happen if 101 the agents vary in their optimality. The link says that if agents vary in their 102 optimality, then there is scope for selection, and the change in the frequency of 103 any gene is given by the covariance between the frequency of that gene in an 104 agent and the agent's relative fitness. The first part of this is intuitive – non-105 zero variance in fitness implies that natural selection can operate; the second 106 part follows from the Price equation with the second term set to zero, described 107 below. A sixth link is discussed by G&G, but we do not treat it separately here 108 as it is a logical consequence of links four and five taken together (as they note 109 in their Appendix.) 110

These links may seem obvious, but as Grafen (2002) points out, that is only 111 112 because many biologists simply take for granted that selection leads to optimization. And in fact, the assumptions that must be made, and the definition 113 of 'fitness' that must be used, in order for the links to be proved are non-trivial 114 matters. For example, when the agents are individuals, the absence of mutation 115 and gametic selection must be assumed to prove the links; and depending on 116 whether the individuals socially interact, different definitions of 'fitness' must 117 be used (Grafen 2006). So the project is highly non-trivial. 118

When the agents are individuals, Grafen speaks of an 'individual as max-119 imising agent' (IMA) analogy, to capture the idea that individuals behave like 120 economically rational agents, attempting to maximise the value of their objec-121 tive function. If all five links hold, the IMA analogy is closely tied to the action 122 of natural selection. It is then legitimate to treat individuals as adaptive units, 123 Grafen argues, and to regard natural selection as acting to optimize each indi-124 vidual's phenotype. But where the links do not hold, there is no justification 125 for employing the concept of individual adaptation. 126

G&G apply a similar logic to groups, by developing a 'group as maximising agent' (GMA) analogy. They study the conditions under which the five links hold, with 'agents' understood as groups. These conditions then determine when talk of group adaptationism is valid, i.e. when it is legitimate to regard groups as adaptive units, and natural selection as acting to optimize the group's phenotype. So for G&G, the validity of group adaptationism thus
depends on whether the selection / optimality links hold, where groups are the
optimizing agents. This yields an understanding of group adaptationism that is
both conceptually clear and mathematically precise.

¹³⁶ 3 The selection / optimality links: 'actual' ver ¹³⁷ sus 'possible' definitions

For the selection / optimality links to be formally proved, they need to be ex-138 pressed mathematically. 'Optimality' is defined as maximization of the objective 139 function; 'scope for selection' and 'potential for positive selection' are expressed 140 in terms of the evolutionary change in what Grafen (1985) calls 'p-scores'. For-141 mally a *p*-score is simply a function from the set of individuals in the population 142 to \mathbb{R} . In the simplest case a *p*-score is an indicator function for a particular al-143 lele, indicating the frequency of the allele within an individual (= 0, 1/2 or 1 144 for diploids); the average of this *p*-score over individuals is then the frequency 145 of the allele in the population. Any weighted sum of such indicator functions 146 also counts as a *p*-score, is why a *p*-score can assume any real value. (These 147 weighted sums represent breeding values of phenotypic traits; see Grafen (1985, 148 2002, 2008) for a full explanation of p-scores.) 149

In Grafen's (2002) discussion of the IMA analogy, and in Grafen (2006), he 150 considers the set of all possible *p*-scores in a population, i.e. all functions from 151 the set of individuals to \mathbb{R} , irrespective of whether these functions indicate the 152 frequency of an allele actually found in the population (or a weighted sum of 153 such functions). So even if two individuals are genotypically identical, there is 154 still some possible p-score for which they differ. Grafen (2002) then defines 'no 155 scope for selection', an expression that occurs in links 1 and 3, as 'no expected 156 change in population-wide average *p*-score, for any possible *p*-score.' Let us call 157 this definition 'no scope for selection (possible)'. 158

Grafen's definition of 'no scope for selection' may seem odd; surely it would 159 be more natural to define it in terms of actual *p*-scores, rather than all possible 160 *p*-scores? An 'actual *p*-score' may be defined as an indicator function for an allele 161 that is actually present in the population (or a weighted sum of such functions). 162 So for a given population, the set of actual *p*-scores is a proper subset of the 163 set of all possible p-scores. If 'no scope for selection' were defined in terms of 164 actual p-scores, it would mean that there whenever there is no expected genetic 165 change in a population, there is no scope for selection, and vice-versa. Let us 166 call this definition 'no scope for selection (actual)'. 167

The biological meaning of the condition 'no scope for selection (actual)' is obvious, but what about 'no scope for selection (possible)'? In effect, the latter condition means that no allele actually present in the population will change in expected frequency *and* that no *neutral* mutations can change in expected frequency. (By contrast, 'no potential for positive selection' concerns the fate of non-neutral mutations.) Conversely, if there *is* 'scope for selection (possible)' in a population, this means that the fitness distribution is such that, if the requisite
genetic variation were present, there would be expected gene frequency change.
So although the condition 'no scope for selection (possible)' seems odd at first
sight, referring as it does to non-actual *p*-scores, it can be given a reasonable
biological interpretation.

Moreover, the 'possible' definition is crucial to Grafen's project. To see why, 179 consider link 5 – which says that if agents vary in optimality then there is 180 scope for selection. Suppose a population of individuals does exhibit variance in 181 optimality (fitness), but is in population-genetic equilibrium. This could be for 182 a number of reasons, e.g. overdominance. For example, suppose that individual 183 fitness depends exclusively on genotype at a single heterotic locus; assume that 184 AA and BB individuals are non-viable, while ABs are viable. So at equilibrium, 185 the individuals do vary in optimality. However, at the locus in question there 186 will be no evolutionary change; and we may assume that at every other locus, 187 all individuals are genotypically identical. So no allele present in the population 188 will change in expected frequency; thus there will be no expected change in any 189 actual p-score. However, there does exist some possible p-score, e.g. whose value 190 is positively correlated with individual fitness, which will change in frequency. 191 So for link 5 to be true in the IMA case, 'no scope for selection' has to be defined 192 with reference to all possible *p*-scores, rather than just actual *p*-scores. 193

It might be argued that the use of 'all possible' p-scores, in the definition of 194 'scope for selection', is unnecessary, as a referee suggests, for the following rea-195 son. In the overdominance example, there are exactly two possibilities: either 196 all individuals are genetically identical at all loci over than the overdominant 197 locus (case A), or this is not so (case B). Both possibilities are consistent with 198 the model assumptions. If we do not know whether case A or case B obtains, 199 then for all we know, there may be an allele actually present in the population 200 which will change in expected frequency. Since we cannot rule this out, in this 201 sense there is 'scope for selection' based solely on change in actual *p*-scores. 202 The problem with this reasoning is that it makes the existence or otherwise 203 of 'scope for selection' dependent on our knowledge, rather than a matter of 204 objective fact. We regards this as undesirable, since the holding of the selection 205 optimality links, and thus the validity of adaptationism, would then become 206 knowledge-relative too. To avoid these untoward consequences, one must allow 207 that there is 'scope for selection' in both cases A and B above, which is precisely 208 what Grafen (2002) achieves by defining 'scope' in terms of all possible p-scores. 209 So the distinction between 'actual' and 'possible' *p*-scores is necessary. 210

In the IMA case, it is easy to see that link 5 is the only link that could not 211 be proved using the weaker 'actual' definition of 'no scope for selection', under 212 the assumptions of no mutation or gametic selection. (The expression 'scope for 213 selection' does not occur in links 2 and 4, while links 1 and 3 must hold on the 214 'actual' definition whenever they hold on the 'possible' definition.) However we 215 show below that in the GMA case, link 5 can hold even on the 'actual' definition 216 of 'no scope for selection', in certain special cases; and moreover, link 5 can fail 217 to hold even on the 'possible' definition, in certain other cases. 218

If we accept the basic logic of the formal Darwinism approach – that adap-

tationism is defined by the five selection / optimality links holding - then the 220 distinction between the 'actual' and 'possible' definitions of 'scope for selection' 221 gives rise to two subtly different forms of adaptationism. It is an open ques-222 tion which is better. Queller and Strassman (2009) have recently argued that 223 whether some entity is a 'unit of adaptation' depends on the extent of actual, 224 not possible, selective processes within that entity. We do not take a stand 225 on this issue here. In what follows we do not endorse either of the two defini-226 tions of 'scope for selection' as objectively correct, but rather explore the logical 227 consequences of both. 228

²²⁹ 4 Groups as Adaptive Units

Gardner and Grafen (2009) consider a model of evolution in a structured population. There are M groups each containing N individuals. Each individual has a genotype, a phenotype, and a reproductive success value. Each group has a 'group genotype', which is an unordered list of the genotypes of its constituent individuals; group genotype determines group phenotype, which determines group reproductive success. As before, a *p*-score is a function from the set of $M \ge N$ individuals to \mathbb{R} .

The evolutionary change in any p-score is described by the change in the 237 average p-score in the population over one generation, which we denote $\Delta \overline{p}$. 238 Gardner and Grafen treat $\Delta \overline{p}$ as a random variable, in order to model uncer-239 tainty, and focus on its expected value. Explicitly incorporating uncertainty 240 allows them to handle many biological complexities; however these are not rele-241 vant for our purposes, so to keep the analysis simple we ignore uncertainty and 242 talk about the actual change. This is strictly for simplicity; the expected change 243 is what really matters, and our results could easily be formulated in such terms. 244 Assuming no gametic selection or mutation, $\Delta \overline{p}$ is given by the simplest form 245 of the Price equation: 246

$$\overline{w}\Delta\overline{p} = Cov_{IxJ}(w_{ij}, p_{ij}) \tag{1}$$

where p_{ij} and w_{ij} are the *p*-score and the reproductive success of the j^{th} individual in the i^{th} group respectively; $I = \{1, \ldots, M\}$ is the set of group indices and $J = \{1, \ldots, N\}$ the set of individual indices; and \overline{w} is average reproductive success in the population.

As is well-known, equation (1) can be expanded into a 'multi-level' format, by partitioning the total covariance between individual *p*-score and individual reproductive success into between-group and within-group components, yielding the result first obtained by Price (1972):

$$\overline{w}\Delta\overline{p} = Cov_I(w_i, p_i) + E_I[Cov_J(p_{ij}, w_{ij})]$$
⁽²⁾

where w_i is the average reproductive success of the i^{th} group, p_i the average p-score of the i^{th} group. The first term on the RHS is the covariance between a group's average p-score and its group reproductive success; the second term is the average, or expectation, across groups of the within-group covariance between individual *p*-score and individual reproductive success. Equation (2) is often regarded as decomposing the total change into components corresponding to the effects of 'group selection' and 'individual selection' respectively. This interpretation is standard in the literature on multi-level selection, though it is not the only way that these contested terms have been defined.

It is a familiar point that substantial within-group selection may undermine 264 group functionality, thus preventing the group from behaving as an adaptive 265 unit (Buss 1987, Maynard Smith & Szathmary 1995, Frank 2003). G&G thus 266 consider two models in which within-group selection is completely absent, which 267 should consitute a 'best-case scenario' for group adaptationism. The first in-268 volves purely clonal groups; the second involves non-clonal groups with full re-269 pression of competition, i.e. no within-group variance in fitness. (By this G&G 270 mean no within-group variance in *expected*, rather than realized, fitness – which 271 means that the existence of reproductive division of labour in a group is fully 272 compatible with zero within-group variance in fitness. This is one place where 273 the distinction between realized and expected fitness, which we are ignoring for 274 simplicity, is important.) 275

If there is no within-group selection on a given p-score, the second RHS term of equation (2) will be zero, in which case it reduces to

$$\overline{w}\Delta\overline{p} = Cov_I(w_i, p_i) \tag{3}$$

Clearly, equation (3) will apply to any *p*-score that shows no within-group variance. So in the clonal groups model, equation (3) will apply to all *actual p*-scores. Similarly, equation (3) will apply whenever there is no within-group variance in fitness, as in the repression of competition model. Both models imply that for each group, the within-group covariance between fitness and *p*-score is zero, and thus the average of this covariance across groups is also zero.

G&G then claim that in both of these models, the links between the GMA 284 analogy and gene frequency change do obtain (with one proviso), so group adap-285 tationism is valid. This is the central positive claim of their paper. The reason 286 the links hold in these models, they claim, is that the assumption of no within-287 group selection renders equation (3) applicable, which in turn allows the five 288 links to be proved, with the objective function taken to be group fitness, i.e. 289 the average fitness of the individuals in the groups. By contrast, when within-290 group selection is not assumed absent, so the full Price equation (2) must be 291 applied, none of the links can be proved, so it is not legitimate to regard groups 292 as adaptive units. 293

The proviso concerns link 4 in the repression of competition model (which says that if all agents behave equally suboptimally, then at least one mutant can spread). This need not be true, G&G argue, because although an improved group phenotype is possible at the suboptimal equilibrium, "there is no guarantee that the corresponding genetic variants will arrange themselves together in groups in such a way as to give rise to the desired group phenotype" (p. 665). In the clonal case this problem doesn't arise, since any group phenotype can ³⁰¹ be produced by a single genetic variant. So they regard talk of group adap-³⁰² tation as fully justifiable in the clonal case, but only partly justifiable in the ³⁰³ repression-of-competition case.

The significance of this consideration is debatable, since a parallel problem 304 arguably applies at the individual level too. In Grafen (2002), where link 4 is 305 proved for individuals, it is simply assumed that any non-resident phenotype 306 can be produced by a genetic variant – even though this may require several 307 simultaneous mutations at different loci. A parallel assumption could be made 308 in the group case, i.e. that any non-resident group phenotype will be produced 309 by mutation, even if this requires several individuals to mutate simultaneously 310 in which case link 4 would be true. It may be that the required assumption 311 is less plausible in the group than the individual case, but this is an empirical 312 matter. Therefore, we are inclined to regard link 4 as equally defensible, in 313 principle, in both the repression-of-competition and clonal models. But nothing 314 in what follows turns on this. 315

³¹⁶ 5 Clonality versus Repression of Competition

Aside from the proviso concerning link 4, G&G treat clonal groups and compet-317 itively repressed groups on a par. However, there is actually a logical difference 318 between them with respect to links 1 and 3. For simplicity we focus on link 1, 319 which to recall says that if all groups are optimal, then there is no scope for 320 selection. Recall the distinction between the 'actual' and 'possible' definitions 321 of 'no scope for selection' from section 3. If we adopt Grafen's original 'possible' 322 definition, it turns out that link 1 is true in the repression of competition model 323 but *not* in the clonal groups model. 324

Repression of competition implies that there is no within-group variance in 325 fitness. (We do not take this condition to *define* repression of competition, 326 for it is possible that within-group fitnesses may be equal anyway. Repression 327 is a causal mechanism for bringing this about.) The absence of within-group 328 variance in fitness can be expressed by $Var_J(w_{ij}) = 0$ for all groups *i*. This 329 implies that for every possible p-score, $Cov_J(w_{ij}, p_{ij}) = 0$ in each group i, which 330 implies that equation (3) above describes the evolutionary dynamics of each p-331 score. Link 1 then follows immediately; since if all groups are optimal then there 332 is no variance in group fitness, so equation (3) tells us that $\Delta p = 0$ for every 333 possible p-score. 334

Now consider clonality. Note firstly that clonal groups cannot be defined as Var_J(p_{ij}) = 0 for all possible *p*-scores and all groups *i*, i.e. no within-group variance in any possible *p*-score in any group. For this condition is logically unsatisfiable, given that the set of possible *p*-scores is the set of all functions from the set of individuals to \mathbb{R} . That groups are clonal means the absence of within-group variance in any *actual p*-score. But there will be many possible *p*-scores that do show within-group variance, even if the groups are clonal.

This means that the condition $Cov_J(w_{ij}, p_{ij}) = 0$ in each group *i* does not hold for every possible *p*-score in the clonal group model, unlike in the repression of competition model (see Appendix 1). Of course, even if that condition does not hold for a given *p*-score, equation (3) could still apply to that *p*-score if the weaker condition $E_I[Cov_J(w_{ij}, p_{ij})] = 0$ holds, i.e. the average over groups of the within-group covariances is zero. However, this latter condition cannot hold true for all *p*-scores, unless within-group fitnesses are equal. (See Appendix 2, Propostion 2, for proof.)

This means that on the 'possible' definition of scope for selection, link 1 only 350 holds in the clonal groups model if there is no variance in within-group fitnesses 351 in any group. Consider a case where the groups are clonal but within-group 352 fitnesses do vary. In this case, it is not true that if all groups are optimal, 353 $\Delta p = 0$ for every possible *p*-score. It will always be possible to find a *p*-score 354 for which the condition $E_I[Cov_J(w_{ij}, p_{ij})] = 0$ does not hold, and for which 355 Δp will be non-zero. So even if all groups are optimal, there will always be 356 scope for selection unless within-group fitnesses are equal in each group. In 357 fact, the absence of within-group variance in fitness turns out to be necessary 358 and sufficient for all the links to hold, as we show in section 7. 359

In a clonal groups scenario, it is of course possible that within-group fitnesses 360 will be equal. This will be so if individual fitness depends only on individual 361 genotype. But this need not be true. There are various reasons why the members 362 of a clonal group may differ in fitness (aside from chance), e.g. they may receive 363 different amounts of social help. It might still be argued that their *expected* 364 fitnesses will be equal, but this depends on how exactly the states of the world, 365 over which the expectation is taken, are defined. In any case, even if it is 366 assumed that clonal group mates have the same expected fitness, in which case 367 link 1 will hold on the 'possible' definition of scope for selection, it is important 368 to realise that it is not clonality but rather the absence of within-group fitness 369 variance that is responsible for the link holding. 370

Since G&G hold that there can be no scope for selection within clonal groups, 371 in virtue of the clonality, it is clear that they are employing the 'actual' definition 372 of 'scope for selection', on which links 1 and 3 do indeed hold for clonal groups. 373 This definition is perfectly reasonable, but as we saw in section 3, adopting 374 it complicates the formal Darwinism approach to individual adaption, as it 375 makes link 5 logically stronger and thus harder to satisfy. In the group case, 376 adopting the 'actual' definition of scope for selection similarly strengthens link 377 5; as a result, repression of within-group competition no longer suffices for link 378 5 to hold, but clonality does. 379

To understand this, consider the following example. A population contains 380 asexual individuals of two genotypes, A and B, living in groups of size N = 4. 381 Groups are competitively repressed, so within each group all individuals have the 382 same fitness. The population contains exactly three types of group: (AAAA), 383 (BBBB) and (AABB); the group fitness function is non-linear, and is such 384 that w(AABB) > w(BBBB) = w(AAAA). (This is a group-level analogue of 385 over-dominance.) As a result, the population is in equilibrium – no gene will 386 change in frequency – but the groups do vary in fitness (optimality). So link 387 5, which says that if the groups vary in fitness then there is scope for selection, 388 need not be true for competitively repressed groups under the 'actual' definition 389

	'Possible' definition	'Actual' definition
Individual	Links 1–5 \checkmark	Links 1–4 \checkmark , link 5 x
Group-Clonality	Links 2,4,5 $\checkmark,$ links 1,3 x *	Links 1–5 \checkmark **
Group–Repression	Links 1–5 \checkmark	Links 1–4 \checkmark , link 5 x

Table 2: Conditions under which the links hold.

* Links 1 and 3 will hold if clonal group mates have identical fitness.

** Link 5 will fail if group fitness does not depend only on group genotype.

³⁹⁰ of scope for selection.

This counterexample to link 5 depends essentially on the groups being non-391 clonal. This is because for there to be a polymorphic equilbrium with fitness 392 differences between groups, it is essential that some groups contain individuals 393 of different genotypes, given that group fitness depends only on group geno-394 type. Therefore, in the clonal groups model, adopting the 'actual' definition of 395 scope for selection does not allow a counterexample to link 5 to be constructed. 396 (Note however that if the assumption that group fitness depends only on group 397 genotype were relaxed, then link 5 would fail even in the clonal case.) 398

The upshot is that depending on whether we use the 'possible' or the 'actual' definition of scope for selection, the selection / optimality links will hold true under different conditions. These differences are summarized in Table 2, for both the individual and group models, under the standard assumptions of no mutation and no gametic selection.

What should we conclude from this? In one respect competitively repressed 404 groups constitute the better case for group adaptationism, but in another respect 405 clonal groups do. If we adopt the 'possible' definition of scope for selection, 406 then repression of competition guarantees that links 1-5 hold, but clonality 407 does not. Some biologists would regard this as welcome result. Queller and 408 Strassman (2009) have argued that a clonal group should not automatically 409 be regarded as a superorganism, if it shows no functional integration and no 410 social interaction among its constituent individuals; see also Ratnieks and Reeve 411 (1992). In a similar vein, Michod (1999) argues that true higher-level individuals 412 (or superorganisms) must possess mechanisms for conflict suppression. By these 413 authors' lights, an analysis of group adaptation that privileges repression of 414 competition is independently desirable. 415

However if we adopt the 'actual' definition of scope for selection, then clonal groups emerge as the better candidate for the superorganism mantle. On this definition, link 5 fails in the repression of competition model but holds in the clonal groups model (so long as group fitness is assumed to be a function of group genotype). This consideration provides a possible basis, over and above the argument given by G&G in relation to link 4, for treating clonality as the 'best case' for group adaptationism.

⁴²³ The dichotomy between clonality and repression of within-group competi-⁴²⁴ tion, as means for unfiying the evolutionary interests of group members, has ⁴²⁵ relevance in relation to 'major evolutionary transitions'. Multi-cellular organisms typically employ both means; their constituent cells are usually genetically
identical, and the fairness of meiosis serves to repress reproductive competition
between the genes within a single genome, in sexual species. Indeed the assumption of fair meiosis, i.e. the absence of gametic selection, is precisely why links
1 and 3 hold true in the individual model of Grafen (2002), on the 'possible'
definition of scope for selection.

We take no stand on whether the 'actual' or 'possible' definition of scope 432 for selection is preferable, nor therefore on whether clonality or repression of 433 competition constitutes the better case of group adaptationism. Our aim has 434 been to explore the logic of formal Darwinism as applied to groups, under both 435 definitions. However in what follows we focus on Grafen's original 'possible' 436 definition, not because we think it is intrinsically superior, but because it allows 437 us to find necessary and sufficient conditions, that are biologically meaningful, 438 for links 1-5 to hold. 439

Group Adaptation versus Fortuitous Group Benefit

In Adaptation and Natural Selection, G.C. Williams (1966) distinguished be-442 tween 'group adaptation' and 'fortuitous group benefit', as part of his celebrated 443 attack on group selectionism. The former refers to a group feature that evolved 444 because it benefits the group, the latter to a group feature that happens to 445 benefit the group but did not evolve for that reason. So on Williams' view, 446 whether a particular feature constitutes a group adaptation depends crucially 447 on its causal history. A clonal group of non-social aphids, or of some marine 448 invertebrate species, would not count as group adaptation by Williams's lights, 449 for the members of such groups engage in no social behaviour, and the groups 450 exhibit little or no functional organization. If some such groups do better than 451 others, this is most likely a side-effect of differences in individual adaptedness. 452

How does Williams' influential concept of group adaptation relate to the 453 concept defined by the formal Darwinism approach of G&G? The concepts are 454 clearly different; G&G hold that group adaptationism applies to any clonal 455 group, while Williams explicitly rules out some clonal groups. From Williams' 456 viewpoint, the five selection / optimality links which G&G take to define group 457 adaptation could hold 'for the wrong reason', i.e. as a side-effect of individual-458 level processes. This would be so in a case in which there is no within-group 459 variation in fitness, and the individuals in each group engage in no social be-460 haviour. Williams would categorize this as fortuitous group benefit, not group 461 adaptation. 462

This difference between G&G's and Williams' concepts may seem puzzling, since Williams' point was precisely that a trait only counts as group adaptation if it has evolved by a process of group-level selection; and G&G define 'group selection' as "that part of gene-frequency change that is responsible for group adaptation" (p.667). So where does the difference stem from?

The answer is that G&G identify 'group selection' with the between-group 468 component of the multi-level Price equation, i.e. the term $Cov_I(w_i, p_i)$ in equa-469 tion (2); while a proponent of Williams' view must reject this definition. As many 470 authors have pointed out, the multi-level Price equation is arguably a flawed 471 way to decompose the total change into components corresponding to distinct 472 levels of selection (Grafen 1984, Nunney 1985, Heisler & Damuth 1987, Good-473 night et al. 1992; Okasha 2004, 2006). The basic problem is that the covariance 474 between group p-score and group fitness may be positive even in the absence 475 of any causal relation between these variables; groups with a high p-score may 476 be fitter, simply because they contain a higher proportion of intrinsically fit 477 individuals, even if there is no group effect on fitness, and no social behaviour. 478 Arguably it is unhelpful to speak of 'group selection' in such a circumstance; 479 individual selection is responsible for the entirety of the evolutionary change. 480 This is a close corollary of Williams' point that group-beneficial features may 481 arise as a side-effect of individual selection. 482

If we accept that group and individual selection should not be identified with 483 the components of the multi-level Price equation, then what decomposition of 484 the evolutionary change should be used to define them? One promising approach 485 is to use 'contextual analysis', a form of multiple regression analysis (cf. Heisler 486 & Damuth 1987). This permits a solution to the problem that besets the Price 487 approach (i.e. the multi-level decomposition in equation (2)), by isolating the 488 effect of a trait on group fitness once individual effects have been stripped away. 489 The total change can still be partitioned into two components, corresponding 490 to the two levels of selection. The crucial difference with the Price approach 491 is that contextual analysis only identifies a component of group selection when 492 there is a 'group effect' on individual fitness. The method is described fully in 493 section 7. 494

G&G discuss contextual analysis, but appear to regard the distinction be-495 tween the Price and contextual approaches as merely semantic. Clearly it is a 496 semantic matter how we use the expressions 'group selection' and 'individual 497 selection', but the question of whether the causal action of natural selection op-498 erates at the individual or group level is non-semantic. We accept G&G's idea 499 that group selection should be defined as the part of gene-frequency change that 500 is responsible for group adaptation, but we show in section 8 that this does not 501 discriminate between the Price and the contextual definitions of group selection. 502 Some biologists might simply reject Williams' distinction outright, and thus 503 reject the idea that the selection / optimality links might hold 'for the wrong 504 reason'. Anyone doing this would naturally accept the Price decomposition, and 505 G&G's analysis. However many authors, ourselves included, regard Williams' 506 distinction between group adaptation and fortuitous group benefit as important. 507 We show in section 9 that accepting this view does not mean abandoning the 508 formal Darwinism approach altogether. 509

⁵¹⁰ 7 Price's Equation versus Contextual Analysis

Contextual analysis treats every individual in the population as having two trait values, an individual *p*-score and the *p*-score of the group it belongs to. The key question is then whether there is an association between fitness and group *p*-score that does *not* result from an association between fitness and individual *p*-score. This is assessed with a linear regression model:

$$w_{ij} = \beta_1 p_{ij} + \beta_2 p_i + e_{ij} \tag{4}$$

where β_1 is the partial regression of individual fitness on individual *p*-score, controlling for group *p*-score; β_2 is the partial regression of individual fitness on group *p*-score, controlling for individual *p*-score; and e_{ij} is the residual whose variance is to be minimized. Therefore β_2 is the change in individual fitness that would result if the group *p*-score of an individual of fixed *p*-score were changed by one unit – it measures the extent to which differences in group *p*-score explain differences in individual fitness, holding individual *p*-score constant.

If β_2 is zero, this means that an individual's fitness depends only on its own p-score, so any covariance between group p-score and fitness is a side-effect of individual selection. Intuitively this means that individual selection is the only force affecting the evolution of the p-score in the population – at least if we follow Grafen (1984) in defining 'individual selection' in terms of an action's 'effects on the actor's number of offspring alone' (p.83-4). This means that for group selection to operate, β_2 must be non-zero.

It is natural to interpret β_1 and β_2 as measures of the direct causal influence of individual *p*-score and group *p*-score, respectively, on individual fitness. However this interpretation is only valid if the true dependence of w_{ij} on p_{ij} and p_i is linear (as for example in a linear public goods game). Of course, even if the true dependence is non-linear, it is possible to apply equation (4); but in that case the partial regression coefficients cannot be construed as measures of direct causal influence.

Using contextual analysis, we can partition the evolutionary change in *p*score into two components, corresponding to individual and group selection as understood here. To do this, we simply substitute equation (4) into (1). After simplifying, this gives:

$$\overline{w}\Delta\overline{p} = \beta_2 Cov_{IxJ}(p_{ij}, p_i) + \beta_1 Var_{IxJ}(p_{ij}) = \beta_2 Var_I(p_i) + \beta_1 Var_{IxJ}(p_{ij})$$
(5)

Equation (5) constitutes an alternative to the Price decomposition given in equation (2), which to recall is:

$$\overline{w}\Delta\overline{p} = Cov_I(w_i, p_i) + E_I[Cov_J(p_{ij}, w_{ij})]$$
⁽²⁾

Note that equations (2) and (5) are both true; but they slice up the total change in different ways. Which equation we favour depends on whether we think 'individual selection' and 'group selection' should be understood as within-group and between-group selection, or as selection on the component of individual fitness that is due to differences in individual *p*-score, and to differences in group *p*-score.

The contextual approach to multi-level selection, enshrined in (5), tallies 549 neatly with Williams' point that 'fortuitous group benefit' and group adap-550 tation are different matters. In cases of fortuitous group benefit, a trait (or 551 *p*-score) that is individually advantageous leads to an incidental benefit for the 552 group; so group p-score will covary positively with the fitness of both indi-553 viduals and groups. But this association goes away if we control for individual 554 *p*-score, as it alone affects individual fitness; therefore β_2 is zero. On the contex-555 tual approach, the evolutionary change is then solely attributable to individual 556 selection, whereas the Price approach wrongly detects a component of group 557 selection. 558

One limitation of the contextual approach is that if a particular p-score 559 shows no variation within groups, then the partial regression coefficients β_1 560 and β_2 are undefined. This is because the absence of within-group variance in 561 p-score means that an individual's p-score and the p-score of their group are 562 perfectly colinear – so it is impossible to compare the difference in fitness of two 563 individuals with the same group p-score but different individual p-scores, and 564 vice-versa. Although equation (5) cannot be applied in such a circumstance, it 565 still makes sense to ask whether there is a direct causal link between individual 566 (or group) p-score and fitness; it is just that the this causal question cannot be 567 answered by purely statistical means. 568

Now recall the GMA analogy, i.e. the selection / optimality links where the 569 agents are groups. It is because G&G find a close relationship between these 570 links holding and the absence of within-group selection, i.e. $Cov_J(p_{ij}, w_{ij}) = 0$ 571 for all groups i, that they regard group adaptationism as intimately related to 572 the Price approach. If one is persuaded by the alternative contextual approach, 573 it is natural to ask what the relation is between the links holding and the absence 574 of individual selection as defined by contextual analysis, i.e. $\beta_1 = 0$ (cf. Foster 575 2009). 576

A first step towards answering this question is to consider the relation between the absence of within-group selection and the absence of individual selection in the contextual sense. Since the Price and contextual partitions slice up the total change differently, one might think that the absence of within-group selection would be logically unrelated to the absence of individual selection in the contextual sense. Surprisingly, it turns out that this is not so.

In Appendix 2 (Proposition 1), we show that the following relation holds. For 583 a particular *p*-score, if there is no within-group selection on that *p*-score, then 584 either $\beta_1 = 0$ or else the *p*-score shows no within-group variance – in which case 585 β_1 and β_2 are undefined. Conversely, if $\beta_1 = 0$, or if the *p*-score shows no within-586 group variance, it follows that there is no change due to within-group selection, 587 i.e. the average across groups of the within-group covariance between p-score 588 and individual fitness is zero. But this does not imply that $Cov_J(p_{ij}, w_{ij}) = 0$ 589 for all groups i. So in short, for a given p-score, "no within-group selection" 590 implies "no individual selection (contextual) or the p-score shows no within-591

no within-group selection	$\forall i \ Cov_J(p_{ij}, w_{ij}) = 0$
no change due to within-group selection	$E_I[Cov_J(p_{ij}, w_{ij})] = 0$
no individual selection in the contextual sense	$\beta_1 = 0$
no within-group variance in fitness	$\forall i \; Var_J(w_{ij}) = 0$
no within-group variance in <i>p</i> -score	$\forall i \; Var_J(p_{ij}) = 0$

Table 3: Conditions on a single p-score

⁵⁹² group variance", but not vice-versa.

One might conclude from this that if the absence of within-group selection characterizes group adaptationism, as G&G hold, then the absence of individual selection in the contextual sense cannot also characterize it. But this does not follow, if we adopt the 'possible' definition of scope for selection, discussed above. For surprisingly, when all possible *p*-scores are considered, the difference in logical strength between the conditions 'no within-group selection' and 'no individual selection (contextual)' disappears. We show this in the next section.

600 8 Main Results

In this section we outline our main results; full proofs are in Appendix 2. We 601 continue to use the basic G&G model of evolution in a structured population 602 outlined in section 3; notation remains unchanged. As before, gametic selection 603 and mutation are assumed absent, and uncertainty is ignored. (This latter 604 restriction could easily be relaxed.) In Table 3, we write formal definitions 605 of the following conditions on a given p-score: "no within-group selection", "no 606 change due to within-group selection", "no individual selection in the contextual 607 sense", and "no within-group variance in fitness", and "no within-group variance 608 in *p*-score". These conditions bear interesting logical relations to one another. 609

Proposition 1. For any given p-score, the following logical implications hold:
 "no within-group variance in fitness"

 \Rightarrow "no within-group selection"

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- \Rightarrow "no change due to within-group selection"
- ⇔ "no individual selection (contextual)" or "no within-group variance in p-score"

Note that the first two of these implications hold in one direction only, but the last is an equivalence.

We now consider all possible *p*-scores, and write formal definitions for the corresponding conditions in Table 4. Importantly, the condition "no withingroup variance in any *p*-score" can never be satisfied, for reasons noted earlier. Similarly, the condition "no individual selection in the contextual sense on any *p*-score" can never be satisfied - because β_1 will be undefined for any *p*-score that shows no within-group variance. Note also that the condition "no within-group variance in fitness" for a single *p*-score, and the corresponding condition on all

no within-group selection on any p -score	$\forall p \; \forall i \; Cov_J(p_{ij}, w_{ij}) = 0$
no change due to within-group selection in any p -score	$\forall p \ E_I[Cov_J(p_{ij}, w_{ij})] = 0$
no individual selection in the contextual sense on any p -score	$\forall p \ \beta_1 = 0$
no within-group variance in fitness for any p -score	$\forall i \; Var_J(w_{ij}) = 0$
no within-group variance in any p -score	$\forall p \; \forall i \; Var_J(p_{ij}) = 0$

Table 4: Conditions on all *p*-scores

⁶²⁵ *p*-scores, are identical; since the variable 'p' does not occur in the expression ⁶²⁶ ' $Var_J(w_{ij})$ '. (In the remainder of this section, 'all *p*-scores' refers to all possible ⁶²⁷ *p*-scores, i.e. all functions from the set of individuals to \mathbb{R} .)

Our main result is that the following logical relations obtain between the conditions on all *p*-scores:

⁶³⁰ Proposition 2.

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⁶³¹ "no within-group variance in fitness"

- \Leftrightarrow "no within-group selection on any p-score"
 - \Leftrightarrow "no change due to within-group selection in any p-score"

 \Leftrightarrow "for each p-score, either no individual selection (contextual) or no within-group variance in that p-score"

Note that each of these implications holds in both directions, i.e. they are 636 equivalences. This is a striking result, given that two of the corresponding 637 implications for a single *p*-score hold only in the left-right direction. To un-638 derstand this, consider the first equivalence, between "no within-group variance 639 in fitness" and "no within-group selection on any p-score". In the left-to-right 640 direction, this is trivial. To see that it holds in right-to-left direction, suppose 641 that fitnesses vary in at least one group. It is then possible to define a p-score 642 which will be subject to selection within that group, simply by assigning 1 to 643 each individual who is at least as fit as the group average, and 0 to every other 644 individual. Therefore, the only way that there can be no within-group selection 645 on any p-score is if there is no within-group variance in fitness. (This is why 646 repression of competition, but not all groups being clonal, is sufficient for link 647 1 to hold.) See Figure 1 for an illustration of this point. 648

The second equivalence, between "no within-group selection on any p-score" and "no change due to within-group selection in any p-score", holds for essentially the same reason. Although any particular p-score can exhibit no change due to within-group selection even if it is subject to within-group selection, if all possible p-scores exhibit no change due to within-group selection this can only be because fitnesses are equal in each group, which implies the absence of within-group selection on any p-score.

Next, consider the relation between the selection / optimality links holding (on the 'possible' definition of scope for selection), and the above conditions. It is easy to see that if there is no within-group selection on any *p*-score, then links 1, 2 and 3 hold, where the 'agents' are groups and the objective function is group fitness (= average individual fitness). (This follows from G&G's parallel





1: Population with within-group variance in fitness

2: Population without within-group variance in fitness

A: All possible p-scores B: p-scores satisfying E_I[Cov_J(w_{ij},p_{ij})]=0 C: p-scores satisfying Cov_J(w_{ij},p_{ij})=0 for all groups *i*

A, B, C

Figure 1: The left-hand box depicts a population in which there is within-group variance in fitness. The set of all possible *p*-scores is A. The set of *p*-scores for which there is no within-group selection is C. The set of *p*-scores for which there is no change due to within-group selection is B. Crucially, for any *p*-score in B but not in C, such as p_1 , one can find another *p*-score in A but not in B, such as p_2 (see Lemma 2 in Appendix 2.) So in a population for which A=B, there can be no *p*-scores in B but not in C, i.e. the sets A, B, and C co-incide. Moreover, the only way in which A can equal B, is if there is no within-group variance in fitness, as in the right-hand box (see Proposition 2).

analysis in relation to all actual p-scores; see their Appendix.) We argued in section 4 that link 4 also holds in the absence of within-group selection on any p-score, and we showed in section 5 that link 5 holds in the same circumstance. Therefore, the absence of within-group selection on any p-score is sufficient for all the links to hold. Our results show that the absence of within-group selection on any p-score is equivalent to no within-group variance in fitness; so the latter condition is also sufficient for the links to hold.

What conditions are necessary for the links to hold? G&G do not explicitly 668 discuss this; they say only that if there is within-group selection on some *p*-score, 669 then the links are "not proven", which is weaker than saying that they do not 670 hold. But the latter is in fact true. If there is within-group selection on some 671 *p*-score, it is easy to show that not all of the five links can be true. In fact, 672 something stronger can be shown, namely that either link 1, 3 or 5 must fail; 673 see Appendix 2. So links 1, 3 and 5 jointly imply the absence of within-group 674 selection on any p-score, which as we have seen is equivalent to the absence of 675 within-group variance in fitness. Thus the latter condition is necessary for links 676 1, 3 and 5 to hold, and is thus necessary for all the links to hold. 677

This result is interesting, since it shows that the five links are not logically independent. For since links 1, 3, and 5 together imply the absence of withingroup variance in fitness, which itself is a sufficient condition for all the links to hold – granting our argument about link 4 – it follows that links 1, 3 and 5 imply links 2 and 4. Therefore, G&G's characterization of group adaptationism,
in terms of all five links holding, could in fact be re-expressed as links 1, 3 and
5 holding. This is not a criticism; to characterize a concept axiomatically one
does not have to use the smallest possible axiom set; some redundancy in the
axioms can be illuminating.

Therefore, granting our argument about link 4, we arrive at the following:

Proposition 3. "links 1, 2, 3, 4 and 5 hold"
⇔ "links 1, 3, and 5 hold"
⇔ "no within-group variance in fitness"
⇔ "no within-group selection on any p-score"
⇔ "no change due to within-group selection in any p-score"
⇔ "for each p-score, either no individual selection (contextual) or no within-group variance in that p-score"

These equivalences explain our claim in the previous section that the Price 695 equation approach to multi-level selection, enshrined in equation (2), has no 696 particular link to group adaptationism, if the latter is defined as the links hold-697 ing. For while it is possible to characterize the five links holding in terms of 698 components of equation (2), as "no change due to within-group selection in any 699 p-score", it is equally possible to characterize their holding in terms of contex-700 tual analysis, by referring only to the parameter β_1 of equation (5). For the 701 condition "for each *p*-score, either no individual selection (contextual) or no 702 within-group variance in that *p*-score" can be re-expressed as "for each *p*-score, 703 either $\beta_1 = 0$ or β_1 is undefined". So the G&G analysis of group adaptation-704 ism, under the 'possible' definition of scope for selection, provides no particular 705 reason to favour the Price over the contextual approach to defining levels of 706 selection. 707

⁷⁰⁸ 9 G. C. Williams strikes back

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We argued above that anyone accepting G.C. Williams' concept of group adap-709 tation should distinguish between the selection / optimality links holding for the 710 'right' and the 'wrong' reason. The links will hold for the 'wrong' reason where 711 there is no group functional integration and no social behaviour among individu-712 als; but the individuals in each group happen to be equally well-adapted. This is 713 an example of fortuitous group benefit, for Williams, and contrasts with genuine 714 group adaptation. This means that Williams' concept of group adaptationism 715 should be defined as 'the links hold + X'. But what is 'X'? 716

Intuitively, 'X' refers to the fact of group-level functional integration, e.g.
the existence of a mechanism for repression of reproductive competition. The
key distinction is between a case where within-group fitnesses are equalized by
some such mechanism, and a case where they just happen to be equal, e.g.
because the groups are clonal and individual fitness depends only on individual
genotype. In the former case the individuals in each group share a common
fate; in the latter case they merely have identical fates.

We introduced contextual analysis, as opposed to the Price approach, precisely to capture Williams' distinction. However the results of the last section show that, on the 'possible' definition of scope for selection, the links holding can be characterized using either the Price or the contextual approaches. So the contextual approach, as outlined above, cannot itself capture the elusive condition 'X'. So how should an advocate of Williams' concept proceed?

A natural suggestion is to modify contextual analysis by introducing a coun-730 terfactual test. Consider a simple example in which an individual's fitness de-731 pends only on a single genetic locus. There are two alleles at the locus, one 732 of which confers a fitness advantage. All individuals in a group have the same 733 allele, but there is between-group variance. So within-group fitnesses are equal, 734 but not because of a group-level effect; so the links hold for the wrong reason. 735 Consider a *p*-score which indicates presence or absence of the superior allele. 736 Since this *p*-score shows no within-group variance, β_1 is undefined. So although 737 individual fitness depends only on individual genotype, contextual analysis can-738 not detect this due to insufficient genetic variation. 739

A solution is to consider what would happen if genetic variation were intro-740 duced within groups. Suppose that one or more individuals had their genotype 741 changed at the locus in question, in a way that leads to within-group variation. 742 This will also change the fitness distribution in the population. Consider the 743 new p-score, denoted p', that indicates presence or absence of the superior al-744 lele in this modified population. Since p' does vary within groups, β'_1 is now 745 defined, so contextual analysis can reveal that individual fitness depends solely 746 on individual genotype, not on group effects, i.e. $\beta'_1 \neq 0$. This shows that in the 747 original population, the absence of within-group variance in fitness, and thus 748 the holding of the links, was not due to a mechanism for repressing reproductive 749 competition, but arose simply because of the absence of within-group variance 750 in the crucial genotype. 751

We can generalize this example into an abstract characterization of what 752 it means for the links to hold 'for the right reason'. Consider all the actual 753 p-scores in the population. Take the subset of the actual p-scores that show 754 no within-group variance, for which β_1 is undefined. (This subset will be non-755 empty in the cases that we are trying to rule out.) For each of these p-scores, we 756 introduce within-group genetic variation in the allele that the *p*-score represents, 757 by modifying the genotypes of one or more individuals. This results in a new 758 set of actual p-scores, to which contextual analysis can be applied again, and for 759 which the β'_1 coefficients must be well-defined. If the links hold for the 'wrong' 760 reason, as in the example above, at least one of these β'_1 coefficients will be 761 non-zero. If they hold for the 'right' reason, each of the coefficients will be zero, 762 indicating that in the original population, the absence of within-group fitness 763 variance did not arise simply because the alleles on which individual fitness 764 depended were fixed in each group, so must have been due to a mechanism for 765 repression of reproductive competition. 766

So Williams' concept of group adaptationism can be defined as 'the links holding for the right reason'. We saw above that 'the links holding' is equivalent to "for each *p*-score, either $\beta_1 = 0$ or β_1 is undefined". The 'right reason' can be

characterized as follows: "for all actual *p*-scores for which β_1 is undefined, $\beta'_1 =$ 770 0". The conjunction of these conditions thus defines group adaptationism à la 771 Williams. This definition ensures that the distinction between group adaptation 772 and fortuitous group benefit is respected. In a case of fortuitous group benefit, 773 the links may hold but will not hold for the right reason, and our counterfactual 774 test will detect this. Where the links hold for the right reason, the covariance 775 between group p-score and group fitness, that appears in equation (2), is not 776 simply a side effect of individual selection, but reflects a direct casual influence 777 of group *p*-score on group fitness. This fits well with Williams' insistence that 778 a group adaptation is a feature of a group that benefits it and that evolved for 779 that reason. 780

The 'right reason' condition may seem unwieldy, referring as it does to what 781 would happen if certain hypothetical changes were introduced into the popula-782 tion. It would be nicer if group adaptationism could be defined without such 783 complications, in terms of actual statistical parameters. However it is not really 784 surprising that this cannot be done. Williams' concept of group adaptation is 785 explicitly causal, and it is a familiar point that causal relations cannot be fully 786 defined in statistical terms. The distinction we have been trying to capture, 787 between the links holding for the right and wrong reasons, is causal, and will 788 usually be detected by contextual analysis, but not always. Where there is in-789 sufficient genetic variation for the regression coefficients of the contextual model 790 to be defined, the distinction can only be captured by considering counterfactual 791 scenarios. 792

Our proposed modification to the definition of group adaptationism - the 793 'right reason' condition – is expressed in terms of contextual analysis. One 794 might think that this provides a reason to favour the contextual over the Price 795 approach to multi-level selection. But in fact, the 'right reason' condition can be 796 characterized using only parameters of the Price equation partition. Recall that 797 any p-score for which β_1 is undefined must show no within-group variance, and 798 vice-versa. We know from Proposition 1 that if a *p*-score does show within-group 799 variance, then $\beta_1 = 0$ is equivalent to there being no change due to within-group 800 selection, i.e. the second term of the Price equation (2) equals zero. Therefore, 801 in the modified scenario, where by definition each new actual *p*-score does show 802 within-group variance, the requirement that $\beta'_1 = 0$ is equivalent to the absence 803 of change due to within-group selection on the *p*-score. Therefore the 'right 804 reason' condition, like the 'links holding' condition, can be equally characterized 805 in terms of the contextual or the Price partitions. 806

Despite this equivalence, the contextual characterization of the 'right reason' 807 condition is more natural. For the condition $\beta'_1 = 0$ has a natural causal inter-808 pretation; it means that the gene in question does not directly affect individual 809 fitness. By contrast, the condition 'no change due to within-group selection' 810 has no natural interpretation. For note that, for a given p-score, this condition 811 is not equivalent to the absence of within-group selection on the p-score. So 812 one cannot capture the 'right reason' condition by requiring that there be no 813 within-group selection in the modified population. 814

10 Conclusion

The idea that groups can be adaptive units is a venerable one in biology, but 816 until Gardner and Grafen's analysis had never received a sufficiently precise 817 formulation. Our approach has been one of critical sympathy with G&G's anal-818 ysis. We have endorsed the essence of the 'formal Darwinism' project – defining 819 adaptationism in terms of links between natural selection and optimality – and 820 followed G&G's lead in applying this methodology to the issue of group adapta-821 tionism. Our aim has been two-fold: firstly to explore the logical consequences 822 of the distinction between the 'actual' and 'possible' definitions of 'scope for 823 selection'; and secondly to see whether G&G's analysis can be reconciled with 824 825 G.C. Williams' influential concept of group adaptation. Our main results are the following. 826

- 827 On the 'actual' definition of 'scope for selection':
- 1. All five selection / optimality links hold for clonal groups (presuming that group fitness is a function of group phenotype alone).
- 2. Link 5 fails for competitively repressed groups.
- On the 'possible' definition of 'scope for selection':
- 3. All five selection / optimality links hold for competitively repressed groups.
- 4. Links 1 and 3 fail for clonal groups (unless there is no within-group variance in expected fitness.)
- 5. The absence of within-group variance in fitness is both necessary and sufficient for the five selection / optimality links to hold.
- 6. Links 1, 3 and 5 are jointly equivalent to links 1, 2, 3, 4 and 5.
- The links holding can equally be characterized in terms of the Price equation or contextual analysis.
- ⁸⁴² In general:

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- 8. Williams' concept of group adaptation can be defined as the links holding
 6. 'for the right reason'.
- 9. The links holding 'for the right reason' can equally be characterized in
 terms of the Price equation or contextual analysis.
- B47 10. Group adaptationism, in Williams' sense, cannot be fully characterized
 without reference to causality.

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A Appendix

A.1 Clonality versus repression of competition

Here we show that on the 'possible' definition of scope for selection, repression
of competition suffices to prove link 1, but clonality does not.

Link 1 can be formally expressed as

$$\forall p \text{ [all groups optimal} \to \Delta \overline{p} = 0] \tag{6}$$

where $\forall p'$ means for all *possible p*-scores.

For a given *p*-score, it follows from equation (2) that $\Delta \overline{p} = 0$ holds if $Var_J(p_{ij}) = 0$ for all groups *i* (case 1), or if $Var_J(w_{ij}) = 0$ for all groups *i* (case 2). Focusing on case 1, we get:

$$Var_J(w_{ij}) = 0$$
 for all groups $i \to (\text{all groups optimal} \to \Delta \overline{p} = 0)$ (7)

Since (7) holds for each *p*-score, we can generalize to:

$$\forall p \; [Var_J(w_{ij}) = 0 \text{ for all groups } i \to (\text{all groups optimal } \to \Delta \overline{p} = 0)]$$
(8)

864 This implies:

 $\forall p \; [Var_J(w_{ij}) = 0 \text{ for all groups } i] \rightarrow \forall p \; [\text{all groups optimal} \rightarrow \Delta \overline{p} = 0] \quad (9)$

Since the condition $Var_J(w_{ij}) = 0$ for all groups *i* does not depend on *p*, equation (9) can be re-written:

$$Var_J(w_{ij}) = 0$$
 for all groups $i \to \forall p$ [all groups optimal $\to \Delta \overline{p} = 0$] (10)

Equation (10) tells us that if all groups are competitively repressed, then link 1 holds on the 'possible' definition of scope for selection.

Now consider case 2. For a given p-score, we have:

$$Var_J(p_{ij}) = 0$$
 for all groups $i \to (\text{all groups optimal} \to \Delta \overline{p} = 0)$ (11)

Just as (7) leads to (9), so (11) leads to:

 $\forall p \ [Var_J(p_{ij}) = 0 \text{ for all groups } i] \rightarrow \forall p \ [all groups optimal \rightarrow \Delta \overline{p} = 0]$ (12)

The LHS of (12) says that no possible *p*-score shows within-group variance, which is impossible. So (12) is vacuously true, but does not say that link 1 holds if all groups are clonal. The latter claim is formally expressed as:

 $\forall \text{ actual } p \left[Var_J(p_{ij}) = 0 \text{ for all groups } i \right] \rightarrow \forall p \text{ [all groups optimal} \rightarrow \Delta \overline{p} = 0]$ (13)

However (13) does not follow from (11).

So under the 'possible' definition of scope for selection, link 1 holds for competitively repressed groups but not for clonal groups.

A.2Main results 877

Proposition 1. For a given p-score: 878

 $\forall i \ Var_J(w_{ij}) = 0$ 879

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 $\Rightarrow \forall i \ Cov_J(p_{ij}, w_{ij}) = 0$ 880

 $\Rightarrow E_I[Cov_J(p_{ij}, w_{ij})] = 0$ $\Leftrightarrow \beta_1 = 0 \text{ or } \forall i \ Var_J(p_{ij}) = 0$

Proof. The first two implications are trivial, so we just prove the last equiva-883 lence. 884

Recall the Price partition: 885

$$\overline{w}\Delta\overline{p} = Cov_I(w_i, p_i) + E_I[Cov_J(p_{ij}, w_{ij})]$$
(2)

And the contextual one : 886

$$\overline{w}\Delta\overline{p} = \beta_2 Var_I(p_i) + \beta_1 Var_{I\mathbf{x}J}(p_{ij}) \tag{13}$$

1. In the right-to-left direction: 887

Suppose that either $\beta_1 = 0$ or $\forall i \ Var_J(p_{ij}) = 0$. 888

• If $\forall i \ Var_J(p_{ij}) = 0$, then both terms of the Price partition are zero, as 889 are both terms of the contextual partition. 890

• If $\beta_1 = 0$, then from the contextual partition we have $\overline{w}\Delta \overline{p} = \beta_2 Var_I(p_i)$. 891 Now consider the group term of the Price equation, $Cov_I(p_i, w_i)$. This 892

always equals $Cov_{I_{XJ}}(p_i, w_{ij})$. Substituting for w_{ij} using regression equa-893

tion $w_{ij} = \beta_1 p_{ij} + \beta_2 p_i + e_{ij}$, we get $Cov_{IxJ}(p_i, w_{ij}) = \beta_2 Cov_I(p_i, p_i) =$ 894

 $\beta_2 Var_I(p_i)$. Therefore $Cov_I(p_i, w_i) = \beta_2 Var_I(p_i) = \overline{w}\Delta \overline{p}$, so $E_I[Cov_J(p_{ij}, w_{ij})] = \overline{w}\Delta \overline{p}$. 895 0 896

- 2. In the left-to-right direction: 897
- Suppose that $E_I[Cov_J(p_{ij}, w_{ij})] = 0$. Therefore, $\overline{w}\Delta\overline{p} = Cov_I(p_i, w_i) = Cov_{I\times J}(p_i, w_{ij})$. 898
- Again, substitute in the regression equation $w_{ij} = \beta_1 p_{ij} + \beta_2 p_i + e_{ij}$, which gives: 899
- $\overline{w}\Delta\overline{p} = Cov_{IxJ}(p_i, w_{ij}) = \beta_1 Var_I(p_i) + \beta_2 Var_I(p_i).$ 900
- But the contextual partition says that: 901
- $\overline{w}\Delta\overline{p} = \beta_2 Var_I(p_i) + \beta_1 Var_{IxI}(p_{ij}).$ 902
- Therefore, $\beta_1 Var_I(p_i) = \beta_1 Var_{I \times J}(p_{ij})$, which implies that either $\beta_1 = 0$, or 903 else $Var_I(p_i) = Var_{I \times J}(p_{ij}).$ 904

If $Var_I(p_i) = Var_{IXJ}(p_{ij})$, then the total variance in *p*-score equals the between-905 group variance, which implies that $\forall i \ Var_J(p_{ij}) = 0$. 906

Lemma 1. $\forall i \ Var_J(w_{ij}) = 0 \Leftrightarrow \forall p \ \forall i \ Cov_J(p_{ij}, w_{ij}) = 0$ 908

Proof. The left-to-right direction is trivial. In the right-to-left direction, suppose 909 that $\forall p \; \forall i \; Cov_J(p_{ij}, w_{ij}) = 0$ but that there is some within-group variance in 910 fitness, i.e. $\exists i \ Var_J(w_{ij}) \neq 0$. Assume without loss of generality that there is 911 only one group q in which fitnesses vary. Now consider the p-score p' defined as 912

follows: $p'_{qj} = 1$ for all individuals j in group g such that $w_{gj} > w_g$, and $p'_{qj} = 0$ 913 for all other individuals in g. By construction, in group g, $Cov_J(p_{qj}, w_{qj}) > 0$, 914 which contradicts the hypothesis. 915 916

Lemma 2. $\forall p \; \forall i \; Cov_J(w_{ij}, p_{ij}) = 0 \Leftrightarrow \forall p \; E_I[Cov_J(w_{ij}, p_{ij})] = 0$ 917

Proof. The left-to-right direction is obvious: the average of components equal to 918 zero is zero. In the right-to-left direction, suppose that $\forall p$ -score $E_I[Cov_J(w_{ij}, p_{ij})] =$ 919 0, but that there exists a p-score p and a group g such that $Cov_J(w_{qj}, p_{qj}) > 0$. 920 As the non-zero covariances must average out, it follows that there exists a 921 group g' such that $Cov_J(w_{q'j}, p_{q'j}) < 0$. 922

Assume without loss of generality that g and g' are the only groups in which 923 the covariances are not zero. Recall that a *p*-score is just a function $IxJ \to \mathbb{R}$. 924 Let a new p-score p' be defined as follows: $\forall i \neq g' \; \forall j \; p'_{ij} = p_{ij}$, and $\forall j \; p'_{a'j} =$ 925 $1 - p_{q'j}$. Informally put, p' is identical to p in general, except in the groups for 926 which the covariance is negative, where they are opposite. 927

So $\forall i \neq g' Cov_J(w_{ij}, p'_{ij}) = Cov_J(w_{ij}, p_{ij}) \geq 0$; but $Cov_J(w_{g'j}, p'_{g'j}) =$ 928 $-Cov_J(w_{g'j}, p_{g'j}) > 0.$ 929

Overall, $\forall i \ Cov_J(w_{ij}, p'_{ij}) \geq 0$, and the inequality is strict for groups g and 930 g'. As a consequence, $E_I[Cov_J(w_{ij}, p'_{ij})] > 0$, which contradicts the hypothesis. 931 932

Proposition 2. 933

 $\forall i \ Var_J(w_{ij}) = 0$ 934 $\Leftrightarrow \forall p \; \forall i \; Cov_J(p_{ij}, w_{ij}) = 0$ 935 $\Leftrightarrow \forall p \ E_I[Cov_J(p_{ij}, w_{ij})] = 0$ 936

 $\Leftrightarrow \forall p \ [\beta_1 = 0 \ or \ \forall i \ Var_J(p_{ij}) = 0]$ 937

Proof. The first equivalence corresponds to Lemma 1; the second equivalence 938 to Lemma 2. The third one is trivially obtained from the final equivalence of 939 Proposition 1 (if an equivalence holds for any given p, then it holds for all p). \Box 940

Proposition 3. 941

- "links 1, 2, 3, 4 and 5 hold" 942 \Leftrightarrow "links 1, 3 and 5 hold" 943
- $\Leftrightarrow \forall i \ Var_J(w_{ij}) = 0$ 944
- 945
- 946

$$\Leftrightarrow \forall p \ [\beta_1 = 0 \ or \ \forall i \ Var_J(p_{ij}) = 0]$$

Proof. The last three equivalences correspond to Proposition 2. Gardner and 948 Grafen (2009, pp. 12–13) have already proved that $\forall p \; \forall i \; Cov_J(p_{ij}, w_{ij}) = 0$ 949 implies links 1-5, which trivially imply links 1, 3 and 5. Thus we only have to 950 show that links 1, 3 and 5 jointly imply one of the last four conditions above. 951

To show that links 1, 3 and 5 jointly imply $\forall i \ Var_J(w_{ij}) = 0$, suppose that 952 there is some group i such that $Var_J(w_{ij}) \neq 0$. There are three possible cases: 953

• all groups are optimal. Then, for any *p*-score, there will be no change due to between-group selection. However, because within-group fitnesses are not identical in group *i*, we can find a *p*-score for which $E_I[Cov_J(p_{ij}, w_{ij})] \neq$ 0, by Lemmas 1 and 2. Therefore there is scope for selection, so link 1 is false.

• all groups are equally suboptimal. By the same reasoning as in the previous case, there is scope for selection, so link 3 is false.

• groups vary in optimality. Since within-group fitnesses are not identical in group *i*, we can find a *p*-score for which $E_I[Cov_J(p_{ij}, w_{ij})] \neq 0$. Therefore, the change in this *p*-score, $\Delta \overline{p}$ is not equal to $Cov_I(p_i, w_i)$, so link 5 is false.

Therefore, if it's false that $\forall i \ Var_J(w_{ij}) = 0$, then either link 1, 3 or 5 is false.

967 **References**

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