

Doğuş Üniversitesi Dergisi, 20 (2) 2019, 89-103

Asset Price Channel: Evidence from Turkey

Varlık Fiyatı Kanalı: Türkiye Örneği

Gülden ŞENGÜN⁽¹⁾, Akmyrat AMANOV⁽²⁾

ABSTRACT: The power and interaction process of the effects of decisions on monetary policy on economic indicators remains uncertain. The asset price channel explains dynamic interactions of monetary policy. In this study, the effectiveness of asset price channel in Turkey's economy for a sample period of 2003Q1-2017Q4 was examined. The ARDL modeling and bounds testing is used to show long-run relationships between variables. According to the findings; there is a long-run equilibrium relationship between the stock prices and the investment expenditures, while there is no long-run relationship between the stock prices and consumption expenditures. To conclude, stock prices have a predictable effect on the investment expenditures, and stock prices may be a good indicator of economic activities.

Keywords: Asset Price Channel, ARDL Model, Monetary Policy

Öz: Para politikası kararlarının ekonomik göstergeler üzerindeki etkilerinin gücü ve etkileşim süreci belirsizliğini korumaktadır. Varlık fiyatı kanalı para politikasının dinamik etkileşimini açıklamaya çalışmaktadır. Bu çalışmada, 2003Q1-2017Q4 örneklem döneminde Türkiye ekonomisinde varlık fiyatı kanalının etkinliği incelenmiştir. Değişkenler arasındaki uzun vadeli ilişkiyi incelemek amacıyla ARDL modeli ve sınır testi yaklaşımı kullanılmıştır. Bulgulara göre; hisse senedi fiyatları ile yatırım harcamaları arasında uzun vadeli bir ilişki bulunurken, hisse senedi fiyatları ile tüketim harcamaları arasında uzun vadeli bir ilişki bulunmamıştır. Sonuç olarak, hisse senedi fiyatları yatırım harcamaları üzerinde tahmin edilebilir bir etkiye sahiptir. Dolayısıyla, hisse senedi fiyatları ekonomik aktivitelerin iyi bir göstergesi olabilir.

Anahtar Kelimeler: Varlık Fiyatı Kanalı, ARDL Modeli, Para Politikası

JEL Classifications: G12, C32, E52

¹ Istanbul University, Department of Economics, Istanbul, Turkey, sengungulden@gmail.com

² Istanbul University, Department of Economics, Istanbul, Turkey, akmyratamanov@gmail.com

Geliş/Received:04-01-2019; Kabul/Accepted: 18-01-2019

1. Introduction

The purpose of economics is to increase social welfare. This aim can be achieved by more efficient use of factors of production or by finding new resources in the long term. Social welfare is ensured by minimizing cyclical fluctuations in the short term. Central banks are targeting price stability and financial stability in order to minimize the cyclical fluctuations. Therefore, the aim of central banks is to try to eliminate the uncertainties of the power and interaction process of the effects of monetary policy decisions. Two approaches are important in the study of monetary actions: (1) a reduced form that explores the importance of money on economic activities, (2) a structural form that examines how monetary policy affects total demand. The monetary transmission mechanism explains how the policies of central bank affect the macroeconomic variables. When the monetary transmission mechanisms are known, the size and timing of the applied policy can be selected in a consistent manner with the specified target. Second, since the links within the monetary transmission mechanisms are obvious, it is likely to be more appropriate to use the indicators when a policy is applied.

Monetary transmission mechanisms are also divided into many parts. One of them is the asset price channel. The asset price channel explains the effects of changes in prices of assets on total demand. Knowing the impact of the stock prices on consumption and investment expenditures will enable policy makers to predict the course of the economy and, if possible, shape their policies accordingly.

Referring to studies conducted in Turkey, it was observed that the relationship between stock price and many macroeconomic variables such as real economic activity, nominal GDP, industrial production growth and economic growth is examined. However, in these studies, relationship between stock prices and consumption-investment expenditures was not examined separately. Examining the relationship separately is important because the multiplier effects of consumption and investment expenditures in the economy are different.

In this study, the effectiveness of asset price channel in Turkey's economy for a sample period of 2003Q1-2017Q4 was examined. Thus, the impacts of changes in the BIST100 index on consumption and investment expenditures was investigated. Firstly, the general information about the asset price channel was introduced to readers. Then some empirical studies related to asset price channel were mentioned. Finally, the ARDL Bounds Test is used to show long-run relationships and causality between variables. Long-term elasticity relationships between variables are revealed.

2. Literature Review

The asset price channel shows the effects of changes in the relative prices of financial and real assets on total expenditure and production. It is based on the assumption that the financial market is perfect, as is the traditional interest channel. The asset price channel can be examined in two categories.

Wealth effect channel is the first asset prices channel (Ando and Modigliani, 1963; Modigliani and Brumberg, 1954). According to Franco Modigliani's the life-cycle hypothesis, the determinant of consumption expenditure is not the current income but the consumer's expected income to achieve throughout life. Life-long resources consist of; human capital, real capital and financial wealth. Stocks in financial wealth constitute a large part of their lifetime resources. For example, in an expansionary monetary policy implementation, there will be more money than the people would like to keep. The public will evaluate the excess by buying stocks. As the demand for stocks increases, stock prices will increase, in this case, people's wealth will increase. This, in turn will increase total demand and total output. We can illustrate the wealth effect channel as follows:

$$M \uparrow \rightarrow Pe \uparrow \rightarrow W \uparrow \rightarrow C \uparrow \rightarrow Y \uparrow \quad (1)$$

Here M represents the money supply, Pe represents the stock prices, W represents the wealth of the people, C represents consumption expenditure and Y represents national income.

The second wealth effect channel is Tobin's q channel (Tobin, 1970). As in the interest rate channel, the relationship between monetary policy and investment expenditures is established through money and bond markets. In this channel, the relationship between monetary policy and investments is explained through stock prices. Changes in stock prices affect the market value of firms and affect the investment decisions. For a better understanding of the channel, let us imagine that the central bank again applies an expansionary monetary policy. In this case, people will have more money than they would like to keep, and people will consider the surplus by buying stocks. This means that Tobin's q (market value of firms/replacement cost of capital) will increase as the increase in stock demand will lead to a rise in stock prices as well. A higher q means that the replacement cost of capital (new plant and equipment acquisition) is lower than the market value of firm. Firms are able to purchase more new plant and equipment capital by issuing fewer stocks. Thus, investment expenditures and total output will increase. The Tobin's q channel can be illustrated schematically as follows:

$$M \uparrow \rightarrow Pe \uparrow \rightarrow q \uparrow \rightarrow I \uparrow \rightarrow Y \uparrow \quad (2)$$

Here M represents the money supply, P_e represents the stock prices, q represents the market value of firms/replacement cost of capital, I represent investment expenditure and Y represents national income.

Many empirical studies have been carried out on the relation among asset prices, stock market performance and economic activity. We can categorize the work done so far into three categories: (1) the effects of asset prices/returns on final targets, (2) policy instruments and other variables that affect asset prices, and (3) the relationship between stock market performance and economic activity. Now, because it covers the scope of work, the focus will be on the effects of asset prices-returns on final targets. Empirical foreign studies are given in Table 1; local studies are given in Table 2.

Table 1. Foreign Studies

Author	Country Period	Method	Findings and Conclusion
Bosworth, Hymans and Modigliani(1975)	U.S.A 1954-1972	Regression analysis	The changes in nominal stock return rates brought similar changes in reel economic activity and the stock market.
Fama (1990) and Schwert (1990)	U.S.A. 1953-1987 1889-1998	Regression analysis	Changes in the real stock return rates are closely related to future economic activities. Unexpected increases in stock price will lead to further increases in the future dividend, while future discount rates will be lower than the projected level. Therefore, rise in stock price is an indicator that the future production will increase.
Mahdavi and Sohrabian (1991)	U.S.A 1960-1989	Granger-causality test	There is an asymmetric causal relationship between the growth of stock prices and the growth of GNP: while the causation is from growth of stock prices to growth of GNP. Even though stock prices are the good indicator for the future macroeconomic events, the impacts of this indicator should not be exaggerated.
Hassapis and Kalyvitis (2002)	G-7 Economies	VAR model	Unanticipated movements in output and real stock prices are found to play a role in the future economic growth and market valuation of capital, and moreover, the responses of growth and real stock returns, after unanticipated shocks in these variables, move in the same direction across countries and data frequencies.
Ludwig and Sløk (2004)	16 OECD countries 1960-2000	Panel Data Analysis	The long-run responsiveness of consumption to permanent changes in stock prices is higher for countries with a market-based financial system than for countries with a bank-based financial system.
Gan, Lee, Yong and Zhang (2006)	New Zealand 1990-2003	Johansen Maximum Likelihood and Granger-causality tests	Although the New Zealand Stock Index is consistently determined by the interest rate, money supply, and real GDP, there is no evidence that this index is a leading indicator for changes in macroeconomic variables. Moreover, since the New Zealand stock market is smaller than other developed countries, the stock market may be very sensitive to macroeconomic factors of its major trading partners.

Husain (2006)	Pakistan 1959-2004	ECM model and Granger-causality test	Regarding the cause and effect relationship, the analysis indicates a one-way causation from the real sector to stock prices implying that the stock market in Pakistan is still not that developed to influence the real sector of the economy.
Naceur and Ghazouani (2007)	11 MENA countries 1979-2003	Dynamic panel model	There is no significant relationship between banking and stock market development and growth. The reason for it must be linked to underdeveloped financial systems in MENA region.
Pilinkus (2010)	Lithuania, Latvia and Estonia 2000-2008	Johansen Cointegration Analysis	It was found that there is a long-term mutual relationship between selected macroeconomic indicators and stock market index in all countries.
Wickremasinghe (2011)	Sri Lanka 1985–2004	Cointegration Analysis	The findings show that there are both the short and long-run causal relationship between stock prices and macroeconomic variables.
Osamwonyi ve Evbayiro-Osagie (2012)	Nigeria 1975–2005	VECM Model	The results indicate that macroeconomic variables influence stock market index in Nigeria.
Aye, Gupta, Modise (2015)	South Africa 1960-2011	TVP-VAR model	The findings show that the impact of a real stock price shocks on consumption is positive. However, there is a negative spillover from the stock market to consumption during the financial crisis.

Table 2. Local Studies

Author	Country Period	Method	Findings and Conclusion
Kaplan (2008)	Turkey 1987-2006	Cointegration test and VAR model	There is a close connection between stock market prices and real economic activity, and there is causation from stock market prices to real economic activity. Therefore, policymakers can depend on the stock market prices as a leading indicator.
Aydemir (2008)	Turkey 1998-2008	VECM model and Granger-causality test	There is a long-run relationship between stocks returns and GDP and private sector consumption expenditures. Moreover, there is a mutual Granger causality. In addition, a one-way causality has been identified from stock returns to private sector investment expenditures.
Örnek (2009)	Turkey 1990-2006	VAR model	Asset prices channel is not efficient and implemented monetary policy has no effect on the stock market. This finding shows that stock market of Turkey is not depth enough.
Altıntaş and Tombak (2011)	Turkey 1987-2008	VAR model and Granger-causality test	There is a positive relation between international reserves, reel exchange rate and economic growth to stock prices, while there is a negative relation from monetary expansion to stock prices.
Balı, Cinel and Günday (2014)	Turkey 2003-2013	Multiple linear regression models	BIST100 is a dependent variable. Inflation rate and interest rate have a negative effect on BIST100, while GDP and Industrial Production Index have a positive effect on it.
Kaya and Uğurlu (2016)	Turkey 1998-2013	Cointegration and VECM model	There is a mutual causation between stock return rates with private sector fixed investment expenditures and GDP. Moreover, there is a strong relation between real economy and stock return rates.
Öztürk (2016)	Turkey 1991-2015	Toda-Yamamoto Granger causality test	There is a uni-directional causality running from the BIST Finance, BIST Services, BIST Industrial Production indices to GDP. Moreover, the BIST100 index does not Granger cause GDP and vice versa.
Coşkun, Kiracı, Muhammed (2016)	Turkey 2005-2015	Granger causality test and impulse response function	The findings show that there is a one-way causality from BIST to Industrial Production Index, export and import. Moreover, it is seen that there is one-way causality from exchange rate to BIST.

When the empirical studies are examined, it is seen that the relationship between stock price and many macroeconomic variables such as real economic activity, nominal GDP, industrial production growth and economic growth is examined. The findings generally indicate that there is a significant relationship between stock prices and macroeconomic variables. However, in the studies, it was not observed that the relationship between stock prices and consumption-investment expenditures was examined separately. Considering that the multiplier effects of consumption and investment expenditures in the economy are different, a distinction between consumption and investment expenditures should be made. This will allow policymakers to obtain more accurate information and implement appropriate policies.

3. Data And Methodology

In this study, the effectiveness of asset price channel was questioned. Thus the causation and parameters among variables have been investigated. (BIST), (TUK), and (YAT) variables were used to set up the model. (BIST) represents the BIST-100 according to the closing price, (TUK) represents final consumption expenditure of resident households and consumption of non-profit institutions serving households, (YAT) represents gross fixed capital formation. The study covers periods of 2003Q1-2017Q4, and the frequency of time series is quarterly. The logarithm of the variables is taken to establish the relation of flexibility among variables. Moreover, there is a structural break which occurs over the period 2008Q1-2010Q3. A dummy variable was constructed, that takes the value one for these observations, and zero everywhere else. Moreover, descriptive statistics of variables is given in Table 3.

Table 3. Descriptive Statistics of Variables

	BIST	TUK	YAT
Mean	55234.80	2.19E+08	99919933
Median	56038.00	1.85E+08	73476783
Maximum	115333.0	5.23E+08	2.68E+08
Minimum	9475.000	66177164	17948774
Standard Deviation	25084.94	1.16E+08	62405879
Observation Number	60	60	60
Source	EVDS Data Central	EVDS Data Central	EVDS Data Central

According to Table 3, an average consumption expenditure of 2.19E+08 and an average investment expenditure of 99919933 was realized during the 2003Q1-2017Q4 period. While the highest consumption and investment expenditures were realized in 2017Q4, the minimum consumption and investment expenditures were realized in 2003Q1. On the other hand, the average value of BIST was 55234.80. The time series is taken from the Central Bank of the Republic of Turkey EVDS data central and each time series covers 60 observations.

In order to find the appropriate model, it was first tested whether the variables contain a unit root. Since the economic model generally has higher autoregressive processes, the ADF unit root test is applied. When this test is applied, the appropriate number of lags included in the model is determined with the help of the Schwarz information criteria. The results are shown in Table 4.

Table 4. Unit Root with Break Test Results for Series

	ADF			
	Level	Level Prob.	1st Difference	1st Dif. Prob.
LNBIST	-4.549	0.1957	-5.901*	<0.01
LNTUK	-5.245*	0.0397	-9.968*	<0.01
LNYAT	-1.586	0.99	-9.105*	<0.01
MacKinnon (1996) one sided p-values.				

* Significant at the 5 % level

According to the results of the unit root with break tests, LNTUK is stationary at normal level while LNBIST and LNYAT become stationary when the first difference of them are taken. ARDL modeling and bounds testing is suitable when the unit root results are inconclusive. Moreover, it is seen that all series are stationary at the 1st difference, and the hypothesis that the data are I(2) is clearly rejected.

4. Findings

In order to see the effects of changes in the BIST100 index on consumption expenditures, the ARDL model was estimated which included the current and past values of the BIST100 index together with the previous values of consumption expenditures, and this model was called the Model TUK. Moreover, the ARDL model was estimated which included the current and past values of the BIST100 index together with the previous values of investment expenditures to see effects of changes in the BIST100 index on investment expenditures, and this model was called the Model YAT. In the specification and estimation of a ARDL model, the BREAK variable was included to take into account the structural breakdown that the global financial crisis has caused, as well as an intercept and trend as regressors. The appropriate number of lags included in the model is determined with the help of the Akaike and Schwarz information criteria. The estimations of Model TUK are given in Table 5, and estimations of Model YAT are shown in Table 6.

When the effects of changes in the BIST 100 index on consumption and investment expenditures were visually examined, it was found that there was a positive correlation between the variables.

Table 5. Estimation of Model TUK

Model TUK ARDL (2,4)			
Variable	Coefficient	Std. Error	Prob.*
LNTUK(-1)	0.165687	0.121286	0.1786
LNTUK(-2)	0.356271	0.104982	0.0014
LNBIST	0.009231	0.022052	0.6775
LNBIST(-1)	0.049471	0.031017	0.1176
LNBIST(-2)	-0.011945	0.031652	0.7076
LNBIST(-3)	0.008547	0.030397	0.7798
LNBIST(-4)	-0.035096	0.018980	0.0709
BREAK	-0.049633	0.009772	0.0000
C	8.731419	1.737596	0.0000
@TREND	0.005187	0.001088	0.0000
R-squared	0.995010	Akaike info criterion	-5.35217
Log likelihood	159.8608	Schwarz criterion	-4.99050
Prob(F-statistic)	0.000000	Durbin-Watson stat	2.14490

Table 6. Estimation of Model YAT

Model YAT ARDL (1,0)			
Variable	Coefficient	Std. Error	Prob.*
LNYAT(-1)	0.673504	0.042458	0.0000
LNBIST	0.095157	0.021052	0.0000
BREAK	-0.091423	0.012836	0.0000
C	4.861663	0.668291	0.0000
@TREND	0.002347	0.000736	0.0024
R-squared	0.990671	Akaike info criterion	-3.93509
Log likelihood	121.0854	Schwarz criterion	-3.75903
Prob(F-statistic)	0.000000	Durbin-Watson stat	2.31807

It is necessary to test whether the actual functional form of the established models and neglected variable problem and/or unnecessary variable problem exist in the models. The accuracy of the specifications of the models was tested with the Ramsey Reset Test.

Table 7. Ramsey Reset Test

Model TUK				Model YAT			
t-statistic	Value	df	Prob.	t-statistic	Value	df	Prob.
	1.631	45	0.109		1.2279	53	0.2249
F-statistic	2.662	(1.45)	0.109	F-statistic	1.5077	(1.53)	0.2249

* Significant at the 5 % level

The null hypothesis of the Ramsey Reset test suggests that the specification of the model is correct. According to the test results, the null hypothesis cannot be rejected, thus the specification of the model is correct.

Another assumption in linear econometric equations is that there is no autocorrelation (relationship) between the error term values. It is important that the error terms in the model are serially independent, otherwise the parameter estimates will be inconsistent. For this reason, whether there is autocorrelation between error terms is examined according to the Correlogram-Q Statistics, and the results are given in Table 8.

Table 8. Correlogram-Q Statistics

Model TUK					Model YAT				
	AC	PAC	Q-Stat	Prob*		AC	PAC	Q-Stat	Prob*
1	0.097	0.097	0.5594	0.455	1	0.188	0.188	2.2017	0.138
2	0.206	0.218	3.1241	0.210	2	0.062	0.027	2.4416	0.295
3	0.046	0.097	3.2543	0.354	3	0.209	0.200	5.2462	0.155
4	0.024	0.093	3.2890	0.511	4	0.064	0.011	5.5153	0.238
5	0.186	0.252	5.4950	0.358	5	0.026	0.050	5.5620	0.351
6	0.032	0.076	5.5601	0.474	6	0.014	0.045	5.5749	0.472
7	0.044	0.089	5.6899	0.576	7	0.052	0.056	5.7632	0.568
8	0.030	0.109	5.7491	0.675	8	0.093	0.140	6.3764	0.605
9	0.121	0.222	6.7676	0.661	9	0.098	0.135	7.0718	0.630
10	0.062	0.111	7.0376	0.722	10	0.059	0.003	7.3286	0.694

11	0.067	0.074	7.3661	0.769	11	0.021	0.012	7.3614	0.769
12	0.032	0.051	7.4410	0.827	12	0.045	0.085	7.5176	0.822

* Significant at the 5 % level

According to this statistics, the null hypothesis suggests that there is no evidence of autocorrelation in the residuals. The results of the statistics show that there is no autocorrelation.

Once these tests have been applied, the examination of the long-term relationship can begin. One of the advantages of the ARDL model is that it allows the application of the Bounds Test. According to this test, the null hypothesis is that there is no long-term relationship between the variables. Bounds Test results are given in Table 9.

Table 9. Bounds Test

Model TUK					Model YAT				
Test Stat	Value	Signif	I(0)	I(1)	Test Stat	Value	Signif	I(0)	I(1)
F-stat.	10.8395	Finite Sample: n=60			F-stat.	22.9969	Finite Sample: n=60		
		10%	4.20	4.69			10%	4.20	4.69
		5%	4.98	5.52			5%	4.98	5.52
		1%	6.78	7.37			1%	6.78	7.37

* Significant at the 5 % level

It is seen that the F-statistic for the Bounds Test is 10.83 and 22.99, and these values clearly exceeds the 5% critical value for the upper bounds. Therefore, the hypothesis of no long-term relationship is strongly rejected. Another advantage of the ARDL model is that it allows estimating long-term coefficients. Long-term coefficients are shown in Table 10.

Table 10. Long Run Coefficients

Model TUK				Model YAT			
Variable	Coef.	S.Error	Prob	Variable	Coef.	S.Error	Prob
LNBIST	0.042	0.026	0.112	LNBIST	0.291	0.0560	0.00
@TREND	0.010	0.0007	0.000	@TREND	0.007	0.001	0.00

$$\text{LNYAT} = 0.2914 \text{ LNBIST} + 0.0072 \text{ Trend}$$

According to the findings, while changes in asset prices affect investment spending, changes in asset prices have no effect on consumption expenditures. Rise of BIST index stimulates investment expenditures. A 10% increase in BIST leads to a 2.9% increase in YAT. Rise in BIST, increases market value of firms; therefore, raise in stock prices affect investment expenditures. Rise in BIST, does not affect consumption expenditures. All variables are both statistically and theoretically significant.

Table 11. Error-correction form

Model TUK		Model YAT	
BIST		BIST	
ECM Coefficients	-0.478042*	ECM Coefficients	-0.326496*
Standard Error	0.082065	Standard Error	0.038600

* Significant at the 5 % level.

In the error-correction model, it is proved that shocks that can occur in the long-run equilibrium can be corrected. The coefficient in the error-correction was negative and statistically significant as expected. This coefficient indicates the rate at which the short-run deviations resulting from the non-stationary series is adjusted in the next period. In Model YAT, the short-run imbalance that occurs in BIST is adjusted approximately in nine months to the long-run equilibrium level.

5. Conclusion

In this study, the effectiveness of asset price channel in Turkey's economy was examined. The relationship among BIST100 index, consumption expenditures and investment expenditures in the Turkish economy for a sample period of 2003Q1-2017Q4 was investigated. The ARDL Bounds Test was applied as an appropriate econometric approach to carry out this study.

According to findings, first, there is a long-run equilibrium relationship between the stock prices and the investment expenditures, while there is no long-run relationship between the stock prices and consumption expenditures. A rise of BIST100 index stimulates investment expenditures, and a 10% increase in BIST100 leads to 2.9% increase in YAT. A rise in BIST100 increases the market value of firms; therefore, raise at stock prices affect investment expenditures. On the other hand, a rise in BIST100 does not affect consumption expenditures. It is considered that stocks may not constitute an important role in the consumers' lifetime resources in Turkey. Therefore, no significant relationship was found between stock prices and consumption. Second, there is a slow adjustment in the investment expenditures when the stock prices change. The short-run imbalance that occurs in BIST100 is adjusted approximately in nine months to the long-run equilibrium level.

According to these findings, stock prices may be a good indicator of economic activities. Stock prices have a predictable effect on the investment expenditures. Further studies should be done on questioning the determinants of stock prices.

6. References

- Altıntaş, H. ve Tombak, F. (2011). Trkiye’de Hisse Senedi Fiyatları ve Makroekonomik Deęişkenler Arasındaki İlişkinin Ekonometrik Analizi: 1987-2008, *Anadolu Üniversitesi Ekonomi Kongresi II*, 1-21.
- Ando, A. and Modigliani, F. (1963). The ‘Life Cycle’ Hypothesis of Saving: Aggregate Implications And Tests, *The American economic review*, 53(1), 55-84.
- Aydemir, O. (2008). Hisse Senedi Getirileri Ve Reel Sektr Arasındaki İlişki: Ampirik Bir Çalıřma, *Afyon Kocatepe University Journal of Economics and Administrative*, 37(2), 37-55.
- Aye, G. C., Gupta, R. and Modise, M. P. (2015). Do Stock Prices Impact Consumption and Interest Rate in South Africa? Evidence From A Time-Varying Vector Autoregressive Model, *Journal of Emerging Market Finance*, 14(2), 176-196.
- Balı, S., Cinel, M. O. ve Gnday, A. H. (2014). Hisse Senedi Fiyatlarını Etkileyen Temel Makroekonomik Faktrlerin Bist 100 Endeksi’ne Etkisinin lçmlenmesi, *OD Sosyal Bilimler Arařtırmaları Dergisi*, 4(9), 46-50.
- Bosworth, B., Hymans, S. and Modigliani, F. (1975). The Stock Market and The Economy, *Brookings Papers on Economic Activity*, 1975(2), 257-300.
- Cořkun, M., Kiracı, K. ve Muhammed, U. (2016). Seçilmiş Makroekonomik Deęişkenlerle Hisse Senedi Fiyatları Arasındaki İlişki: Trkiye zerine Ampirik Bir İnceleme. *Finans Politik & Ekonomik Yorumlar*, 53(616), 61-74.
- Fama, E. F. (1990). Stock Returns, Expected Returns, and Real Activity. *The Journal of Finance*, 45(4), 1089-1108.
- Gan, C., Lee, M., Yong, H. H. A. and Zhang, J. (2006). Macroeconomic Variables and Stock Market Interactions: New Zealand Evidence. *Investment Management and Financial Innovations*, 3(4), 89-101.
- Hassapis, C. and Kalyvitis, S. (2002). Investigating The Links Between Growth and Real Stock Price Changes With Empirical Evidence From The G-7 Economies. *The Quarterly Review of Economics and Finance*, 42(3), 543-575.
- Husain, F. (2006). Stock Prices, Real Sector and The Causal Analysis: The Case of Pakistan, *Journal of Management and Social Sciences*, 2(2), 179-185.
- Kaplan, M. (2008). The Impact of Stock Market on Real Economic Activity: Evidence From Turkey. *Journal of Applied Sciences*, 8(2), 374-378.
- Kaya, E. ve Uęurlu, S. (2016). Seçili Bazı Makroekonomik Deęişkenler ve Hisse Senedi Piyasası Arasındaki Dinamik Etkileşim: Bist 100 iin Ekonometrik Bir Yaklaşım, *Pamukkale University Journal of Social Sciences Institute/Pamukkale Üniversitesi Sosyal Bilimler Enstits Dergisi*, (23), 1-13.
- Ludwig, A. and Slk, T. (2004). The Relationship Between Stock Prices, House Prices and Consumption in OECD Countries. *Topics in Macroeconomics*, 4(1).

- Mahdavi, S. and Sohrabian, A. (1991). The Link Between The Rate of Growth of Stock Prices And The Rate of Growth of GNP in The United States: A Granger Causality Test. *The American Economist*, 35(2), 41-48.
- Modigliani, F. and Brumberg, R. (1954). Utility Analysis and The Consumption Function: An Interpretation of Cross-Section Data, *The MIT Press*, 1, 388-436.
- Naceur, S. B. and Ghazouani, S. (2007). Stock Markets, Banks, and Economic Growth: Empirical Evidence From The MENA. Region, *Research in International Business and Finance*, 21(2), 297-315.
- Osamwonyi, I. O. and Evbayiro-Osagie, E. I. (2012). The Relationship Between Macroeconomic Variables and Stock Market Index in Nigeria, *Journal of Economics*, 3(1), 55-63.
- Örnek, İ. (2009). Türkiye’de Parasal Aktarım Mekanizması Kanallarının İşleyişi. *Maliye Dergisi*, 156(1), 104-125.
- Öztürk, F. (2016). Türkiye’de Hisse Senedi Fiyatları ve GSYİH Arasındaki İlişkinin Analizi, *Uluslararası Yönetim İktisat ve İşletme Dergisi*, 30, 86-94.
- Pilinkus, D. (2010). Macroeconomic Indicators and Their Impact On Stock Market Performance in the Short And Long Run: The Case of The Baltic States, *Technological and Economic Development of Economy*, 16(2), 291-304.
- Schwert, G. W. (1990). Stock Returns and Real Activity: A Century of Evidence, *The Journal of Finance*, 45(4), 1237-1257.
- Tobin, J. (1970). Money and Income: Post Hoc Ergo Propter Hoc?, *The Quarterly Journal of Economics*, 84(2), 301-317.
- Wickremasinghe, G. (2011). The Sri Lankan Stock Market and The Macroeconomy: An Empirical Investigation, *Studies in Economics and Finance*, 28(3), 179-195.