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Inferring Amazon leaf demography from satellite observations of leaf area index

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Abstract. Seasonal and year-to-year variations in leaf cover imprint significant spatial and temporal variability on biogeochemical cycles, and affect land-surface properties related to climate. We develop a demographic model of leaf phenology based on the hypothesis that trees seek an optimal leaf area index (LAI) as a function of available light and soil water, and fit it to spaceborne observations of LAI over the Amazon basin, 2001–2005. We find the model reproduces the spatial and temporal LAI distribution whilst also predicting geographic variation in leaf age from the basin centre $(2.1 \pm 0.2 \text{ years})$, through to the lowest values over the deciduous eastern and southern Amazon (6 ± 2 months). The model explains the observed increase in LAI during the dry season as a net addition of leaves in response to increased solar radiation. We anticipate our work to be a starting point from which to develop better descriptions of leaf phenology to incorporate into more sophisticated earth system models.

1 Introduction

Seasonal and year-to-year variations in leaf area imprint significant spatial and temporal variability on biogeochemical cycles and affect land-surface properties related to climate (Hayden, 1998). For example, the transfer of water from the soil to the atmosphere is mostly via leaves through evapotranspiration, subsequently affecting humidity, air temperature and rainfall (Wilson and Baldocchi, 2000). Similarly, carbon enters vegetated ecosystems through carbon fixation via photosynthesis (White et al., 1999).

Over temperate regions leaf phenology is known to be driven by changes in day length and temperature (Schwartz, 1999), although the relative importance of these determining factors and how they might change with climate is poorly understood (Korner and Basler, 2010).

However, the majority of the world's forests retain leaves year round: boreal forests which are dominated by evergreen needle-leaf trees and often mixed with deciduous broadleaf and needle-leaf species; and mesic tropical forests, dominated by evergreen broadleaf species, which are responsible for the majority of terrestrial carbon fixation (Malhi and Grace, 2000). Observed LAI over these evergreen forests, particularly over the tropics, still show seasonal and yearto-year variations (Myneni et al., 2007), but we lack knowledge about the magnitude, geography, timing, and the processes driving such variation, partly reflecting the difficulty of taking year-round measurements. Consequently many modelling studies assume that tropical leaf area is constant (Cramer et al., 2001; Arora and Boer, 2005).

Space-borne observations of LAI offer the best opportunity to develop a quantitative model of large-scale controls of leaf area by virtue of their frequency and global coverage. We focus our study on the Amazon basin $(10^{\circ} \text{ N}-10^{\circ} \text{ S}, 80^{\circ} \text{ W}-50^{\circ} \text{ W})$. The vegetation in the region is mainly semideciduous or evergreen tropical forest, but the species composition varies widely due to the differences in soil type and altitude across the basin (Sombroek et al., 2000).

The Amazon basin experiences wet and dry seasons, with the dry season generally running from June–September, with longer and drier periods in the south-eastern regions (Sombroek, 2001). Fig. 1 shows that the timing of low precipitation coincides with an increase in direct radiation mainly due to a decrease in cloud cover. Levels of diffuse radiation are comparatively constant throughout the year.



Fig. 1. From top to bottom: MODIS leaf area index ($m^2 m^{-2}$), WorldClim precipitation (mm) (Hijmans et al., 2005), GEOS-4 direct and diffuse photosynthetically active radiation, PAR (Wm^{-2}) across the study region. The dry season, with precipitation levels of under 100 mm, runs generally from July–September, period which coincides with an increase in direct PAR due to a decrease in cloud cover. We can observe that the LAI also peaks during this period, reaching its lowest levels in the wet season.

Ground-based studies have reported an increase in leaf litterfall during the dry period (Malhado et al., 2009; Chave et al., 2010), but without simultaneous measurements of leaf gain we cannot determine whether the increased litterfall represents a net loss of leaves. Studies using space-borne vegetation data (Myneni et al., 2007; Huete et al., 2006) have reported an increase in greenness during the dry season over the Amazon, even during severe droughts (Saleska et al., 2007), but these drought observations have been disputed (Samanta et al., 2010; Doughty and Goulden, 2008). These observations are consistent with indirect evidence from the seasonal cycle of satellite-observed emission of biogenic trace gases (Barkley et al., 2009). The dry-season increase in leaf area could be explained by soil moisture dynamics: water is available all year round in the deeper soil layers (Harper et al., 2010), which can be readily accessed by the large rooting depths of Amazonian vegetation (Nepstad et al., 1994; Jipp et al., 1998). Under these circumstances, we expect that light availability is the primary controlling factor for determining changes in leaf area (Wright and Vanschaik, 1994). This implies that trees will carry more leaves in the dry season when direct radiation is greater. To test this idea and to enable predictive modelling of Amazon leaf phenology, we develop a simple leaf phenology model for the Amazon tropical forest. Sect. 3 describes this model, which we fit to MODIS LAI data (Sect. 2) to obtain parameter values for the Amazon basin. We discuss our results in Sect. 4 and demonstrate how, in principle, our predictions of leaf area and age distribution could impact carbon assimilation using a simple carbon model. We conclude our paper in Sect. 5.

2 Data sets used

2.1 MODIS LAI data

We use leaf area index (LAI) data obtained from the MODIS (Moderate Resolution Imaging Spectroradiometer) instrument aboard the NASA Terra platform. The LAI/fPAR (fraction of absorbed photosynthetically active radiation) product collection 5 (MOD15A2) is available globally at a spatial resolution of 1 km every 8 days for the period 2000–present and has been downloaded from https://lpdaac.usgs.gov/. The 8 day temporal resolution is a result of compositing, i.e. assigning the best value for the 8 day period based on maximum fPAR. The data set is split into tiles (10° latitude by 10° longitude at the equator), which cover northern South America and include the Amazon basin. We use tiles h10-12v08 and h10-12v09 for the years 2001 to 2005.

The LAI is calculated using a radiative transfer algorithm (the main algorithm), which uses the red (648 nm) and nearinfrared (858 nm) bands. The algorithm uses biome-specific vegetation structure and height, leaf type and soil brightness to obtain LAI values (Yang et al., 2006; Knyazikhin et al., 1997, 1999). In conditions where the main algorithm cannot be applied, a back-up algorithm is used, in which case the LAI value is calculated using an empirical relationship between Normalised Difference Vegetation Index (NDVI) and LAI. The data quality is affected by the presence of cloud, atmospheric aerosol loading and snow cover. Snow cover is not an issue for our study region, but cloud cover can affect the quality of the data, especially during the wet season, while aerosols from biomass burning can interfere with measurements during the dry season. The quality flags provided along with the LAI data offer information on overall data quality, the algorithm used, cloud cover and aerosol presence. Ground validation studies (Cohen et al., 2006) have shown that the back-up algorithm is often unreliable and our analysis of the data over the Amazon region shows that values assigned by the backup algorithm are often unrealistically low, leading to large week-to-week swings in LAI. As a consequence, we remove any LAI values calculated using the back-up algorithm prior to spatial averaging.

LAI retrievals of vegetation often have saturation problems in that LAI becomes insensitive to changes in reflectance. Studies have shown that this was an issue for high-biomass areas for collections 3 and 4, but this has been considerably improved for collection 5 (Yang et al., 2006). Ground based values of LAI in the Amazon basin range from $3.5-6 \text{ m}^2 \text{ m}^{-2}$ (Malhado et al., 2009; Aragão et al., 2005; Roberts et al., 1996; Meir et al., 2000), with values of up to $10 \text{ m}^2 \text{ m}^{-2}$ registered by some studies (Doughty and Goulden, 2008), with differences arising from the different measurement methods. MODIS LAI values are in the range $2-6 \text{ m}^2 \text{ m}^{-2}$, which provides us with some confidence that there are no major saturation problems. When relating leaf reflectance measurements to seasonal cycles we must take into account the fact that changes in observed reflectance can also be caused by changes in the number of young leaves, as these have different reflectance properties. This has been advanced as an explanation for the observed seasonal swings over the Amazon basin by several studies (Doughty and Goulden, 2008; Asner and Alencar, 2010; Aragão et al., 2009; Brando et al., 2010). However, the observed seasonal changes in LAI are too large to be attributed to a flush of new leaves only (Samanta et al., 2012).

The study region includes lowland tropical forests, alpine forests, savannas and grasslands. As our study is focused on forests, we use the MODIS landcover product MOD12Q1 to filter non-forested pixels. We use the provided IGBP classification scheme and have retained only pixels in classes 1– 5, evergreen needleleaf forest, deciduous needleleaf forests, deciduous broadleaf forest, evergreen broadleaf forest and mixed forest. As there is no way to distinguish between lowland and alpine forests we include both in our analysis.

We reproject the LAI data to an orthogonal projection and average it to the resolution of the GEOS-4 PAR data (2° latitude and 2.5° longitude) and subsequently fit our model at this resolution.

2.2 Radiation data

We use photosynthetic active radiation (PAR) fields from assimilated meteorological data products of the Goddard Earth Observing System (GEOS-4) based at the NASA Global Modeling and Assimilation Office (GMAO) (Bey et al., 2001). The temporal resolution of this data is 1 day and the spatial resolution is $2^{\circ} \times 2.5^{\circ}$.

2.3 Soil moisture data

We use the volumetric soil moisture for 10–200 cm depth from the NCAR/NCEP reanalysis daily average surface flux data set (http://www.esrl.noaa.gov/psd/data/gridded/ data.ncep.reanalysis.surfaceflux.html) (Kalnay et al., 1996). The data is available at global scales at daily timesteps for the period 1948 to present on a Gaussian grid. We reproject the data onto the orthogonal GEOS-4 grid for model fitting.

3 Leaf phenology model

The central assumption of our model is that trees adjust their gains and losses of leaves in order to try to achieve, at any given time, an optimal LAI, which we refer to as the target LAI, LAI_{targ}. The value of LAI_{targ} is determined as the minimum of a light-limited target, LAI^{light}_{targ}, and a water-limited target, LAI^{water}_{targ}. We define LAI^{light}_{targ} such that the bottom layer of leaves receives just enough light to return a positive carbon balance, i.e., receives light at the light compensation

Table 1. Model parameters

Symbol	Units	Description
C_{direct} C_{diffuse} p $gain_{\text{max}}$ a_{crit} μ_0	Wm ⁻² Wm ⁻² days m ² m ⁻² years years ⁻¹	Light compensation point for direct PAR (Eq. 1, main paper) Light compensation point for diffuse PAR (Eq. 1, main paper) Length of time window for average PAR Maximum gain rate for new leaves Age after which leaves start ageing Background decay constant
μ_1	years '	Age-related decay constant

point *C* as derived from Beer's law:

$$LAI_{targ}^{light} = -\frac{1}{\alpha} \ln(\frac{C}{I_0}), \qquad (1)$$

where I_0 is the incoming PAR at the top of the canopy and α is the light attenuation coefficient applying to Beer's Law, which we have fixed to 0.5 (Appendix A1). To recognise the potentially important difference between direct and diffuse light we apply Eq. 1 separately for both direct and diffuse PAR, to determine their respective compensation points (C_{direct} and C_{diffuse}), and then keep the minimum of the two values. For both diffuse and direct PAR we assume that, in order to avoid sub-optimal responses to very short-term variation in light, trees calculate the target LAI_{targ}^{light} using an effective I_0 , defined as the average over the previous p days. We drive Eq. 1 with GEOS-4 reanalysis estimates of incoming PAR (Sect. 2.2). We define the water-limited target as $LAI_{targ}^{water} = \beta_1 + \beta_2 W_s$, where W_s is soil moisture (from the NCAR/NCEP Reanalysis, Sect. 2.3) and β_1 and β_2 are empirical coefficients (Appendix A2) We assume leaf demography (the gain and loss of leaves of different ages) is determined by the factor limiting LAItarg. If the current LAI is below LAI_{targ} trees add new leaves of age a = 0 at a maximum rate gainmax to reach LAItarg. If LAItarg is water limited and the current LAI is above LAItarg, to avoid excessive water loss or overheating leaves, trees lose leaves, beginning with the oldest leaves, until they achieve LAI_{targ}. When LAI_{targ} is light limited and LAI is above LAItarg, trees add no new leaves, but do not actively drop leaves. In all of the above cases, leaves are subject to continuous leaf loss according to a mortality rate μ due to leaf ageing that depends only on leaf age a. We define a minimum age, a_{crit} (years), below which we only consider the background loss, e.g. herbivory, branch loss, so that the mortality is $\mu = \exp^{-\mu_0}$. Leaves older than $a_{\rm crit}$ are lost at a faster rate $\mu = exp^{-\mu_1}$ which is caused by leaf ageing. In order to calculate the age-dependent mortality we introduce the concept of leaf cohorts, defined as a group of leaves of the same age. For each cohort $LAI(a, \mathbf{x}, t)$ we apply the mortality rate as:

$$LAI(a, \mathbf{x}, t) = \mu(a)LAI(a - 1, \mathbf{x}, t - 1),$$
⁽²⁾

with the mortality rate μ defined as above. The overall LAI at each time step, LAI(**x**, *t*) is the sum of all leaf cohorts.

Overall the rate of change of LAI at each location \mathbf{x} and time *t* (Fig. 2) is:

$$\frac{a}{dt}(\text{LAI}(\mathbf{x},t)) = P(I_0(\mathbf{x},t),\text{LAI}(\mathbf{x},t-1)) - -L(\text{LAI}(\mathbf{x},t),W_s(\mathbf{x},t)), \quad (3)$$

where $P(I_0(\mathbf{x}, t), \text{LAI}(\mathbf{x}, t-1))$ denotes production processes and $L(\text{LAI}(\mathbf{x}, t)W_s(\mathbf{x}, t))$, refers to loss processes due to both the age-related mortality rate and active leaf dropping due to water stress. When integrated over time *t*, Eq. 3 provides, for each location \mathbf{x} and time *t*, a predicted LAI (Appendix B1) given environmental drivers (direct and diffuse PAR, and available soil moisture), and given the value of 9 parameters specific to location \mathbf{x} : C_{direct} , C_{diffuse} , p, $gain_{\max}$, a_{crit} , μ_0 , μ_1 , β_1 , β_2 (Table 1). Initial parameter estimates (not shown) estimated that $\text{LAI}_{\text{targ}}^{\text{light}} < \text{LAI}_{\text{targ}}^{\text{water}}$ for nearly all locations at nearly all times, such that the fit to data was not statistically improved by considering water limitation of leaf area so that we can set $\text{LAI}_{\text{targ}} = \text{LAI}_{\text{targ}}^{\text{light}}$; consequently we do not discuss soil moisture further.

We simultaneously fit the above parameters using a Bayesian approach (Appendix B) over our study region with a spatial resolution of 2° (latitude) $\times 2.5^{\circ}$ (longitude) to collection 5 of the LAI data from the MODIS satellite instrument, which was spatially averaged to this resolution (Sect. 2.1).

4 Results

Fig. 3 shows that the model reproduces the spatial distribution of mean LAI (Pearson correlation coefficient $r^2 = 0.9$), capturing the high values (up to $4.8 \pm 0.1 \text{ m}^2 \text{ m}^{-2}$) over the central and southern Amazon basin and lower values $(4.0 \pm 0.2 \text{ m}^2 \text{ m}^{-2})$ over the Eastern regions. More importantly, the model reproduces the broad spatial distribution of LAI amplitude, defined as the difference between the maximum and minimum monthly LAI, with a statistically significant correlation coefficient of $r^2 = 0.46$. This result supports



Fig. 2. How to calculate the predicted leaf age distribution and predicted total LAI for a model driven only by light. Note that the calculation begins with the leaf age distribution from a previous time $t - \delta t$ at the same location, which is updated to account for background leaf mortality, and the addition of new (age = 0) leaves as driven by the LAI target, producing a predicted leaf age distribution for time *t*.



Fig. 3. MODIS (top) and model (bottom) leaf area index (LAI) over the Amazon $(10^{\circ}N-10^{\circ}S, 80^{\circ}W-50^{\circ}W)$, 2001–2005 averaged over a 1×1 km grid. (a) mean LAI (m² m⁻²), (b) mean annual amplitude of LAI (m² m⁻²), and (c) mean timing of peak LAI (day of year).

our model structure because, unlike the mean LAI, the LAI amplitude is highly constrained by model assumptions; the maximum LAI amplitude is determined by the amplitude of incoming PAR.

Figure 4 shows that the model generally has a negative bias with respect to amplitude, which we attribute at least in part to measurement noise, with mean MODIS (model) LAI amplitude of 1.5 ± 0.4 (0.7 ± 0.4) m² m⁻², but the MODIS value falls within the confidence intervals of the model (Fig. 5). Similarly, the model reproduces the seasonal timing of LAI variation (Fig. 4), which is also highly constrained by the model structure, as the greatest target LAI occurs at the time of peak incoming PAR. We find that our model generally describes between 20–80 % (median of 31 %) of the observed temporal variability of LAI at any one $2^{\circ} \times 2.5^{\circ}$ grid cell.

Fig. 6 shows posterior model parameters, which provide further insight into the underlying processes that determine observed variations in LAI. The spatial variations of the parameters are a reflection of not only the seasonality but also of species composition, soil type or nutrient availability. The two compensation points, C_{direct} and C_{diffuse} , can be interpreted as a measure of the shade adaptation in trees, with a lower compensation point indicating leaves that are adapted for lower light conditions. We estimate that C_{direct} is lower in the south of the Amazon, with values of 1.5 compared to 5 Wm^{-2} elsewhere. In contrast, C_{diffuse} , which effectively limits the overall compensation point during the dry season resulting in a lower LAI amplitude, is more homogeneous across the basin with mean values of 0.23 Wm^{-2} . Our values for the compensation points are broadly consistent with ground-based measurements (Riddoch et al., 1991), providing further support for our methodology.

The delay *p* represents the time required for the vegetation to respond to changes in PAR. We find that *p* is generally 14 days for most of the basin. The exception is over the northwestern region, where p > 1 month, suggesting that vegetation over this region is slower to respond to changes in PAR. The maximum gain of LAI is typically around $1 \text{ m}^2 \text{ m}^{-2} \text{ month}^{-1}$, with the highest values (up to $2.2 \text{ m}^2 \text{ m}^{-2} \text{ month}^{-1}$) over the eastern, drier regions corresponding to an area with a higher LAI amplitude and a more pronounced seasonal cycle.



Fig. 4. LAI time series in (a) the eastern Amazon (8° N 62.5° W), (b) the semi-deciduous Amazon (0° N 72.5° W), and (c) the evergreen central Amazon basin (4° N 67.5° W), as predicted by the model (black line) and MODIS LAI data (blue line). Gray shaded area represents 95% confidence intervals. Blue bands represent aproximate dry seasons.



Fig. 5. Model uncertainty for mean LAI and annual amplitude (the difference between the maximum and minimum monthly LAI). We used samples drawn from the posterior distribution to calculate model LAI values and then obtain posterior means and confidence intervals. Here, the uncertainty is represented as the difference between the upper and lower bounds of the 95 % conidence intervals using parameter posterior distributions.

To help interpret our estimated leaf loss parameters, determined by the amplitude of the observed LAI seasonal cycle, we calculate a leaf lifespan τ_{95} as the time at which 95 % of the leaves from a cohort have dropped, based on the exponential decay rates μ_0 and μ_1 (Eq. 2):

$$\tau_{95} = -\frac{\ln 0.05}{\mu_1} - \frac{\mu_0}{\mu_1} a_{\rm crit} + a_{\rm crit},\tag{4}$$

with variables as defined in Table 1. Figure 6 shows that τ_{95} is longest in the middle of the Amazon basin, with values of around 2.1 ± 0.2 years, and lower in the Southern and Eastern regions (1.5 ± 0.7 years), where the vegetation has a larger deciduous component.

These lifetimes are consistent with sparse ground-based studies over the same region, which report values between 2

months and 6.4 years (Reich et al., 2004) and in other tropical forests (up to 26 months (Sharpe, 1997; Osada et al., 2001)).

To obtain an estimate of parameter uncertainty, we use the posterior distribution resulting from the fitting algorithm to calculate 95% confidence intervals (Fig. 7). Most parameters are well constrained, with confidence intervals of 0.1 (± 0.07) of the posterior mean for most parameters. The exceptions are two of the leaf mortality parameters (Fig. 7f and g), with confidence intervals of 0.8 (± 0.2) for the base mortality rates, μ_0 and 0.5 (± 0.1) for the age related rate, μ_1 . This can indicate a trade-off between the two parameters as they both contribute to determining the overall leaf lifespan. Also, the diffuse compensation point C_{diffuse} (Fig. 7b) is less well constrained in the north-eastern regions.



Fig. 6. Mean posterior parameters that describe leaf gain and loss: (a) direct PAR compensation point, C_{direct} (W m⁻²); (b) delay in vegetation response to changes in PAR, p (days); (c) maximum number of leaves that can be added over a month, gain_{max} (m² m⁻² month⁻¹); and (d) mean cohort leaf lifetime τ_{95} defined as the time at which 95% of the leaves from a cohort have dropped (years).

Our model also allows us to estimate the leaf age distribution at any point over the basin, something that would be extremely difficult to do using traditional means. Fig. 8 shows that leaves in regions with a high leaf turnover rate are generally younger than one year, with a high proportion of very young leaves (less than 6 months) with an approximately exponential leaf age distribution which shows pronounced seasonality between wet and dry seasons. In contrast, over the evergreen areas of the central Amazon basin we estimate a higher proportion of leaves older than 1 year and a leaf age distribution with a less pronounced seasonality. In the more deciduous regions in the southern basin, we find distinct leaf cohorts originating from past growing seasons.

To provide an example of the potential impact of our new estimates of leaf age distribution on large-scale calculations of biogeochemistry, we incorporated this information into a simple model of carbon assimilation. We present three scenarios: (1) using a constant LAI and constant leaf age distribution throughout the year, (2) using the predicted LAI with a constant leaf age distribution and (3) using the predicted LAI and leaf age distribution. Fig. 9 shows that the seasonality of the carbon flux is driven mainly by the incoming PAR and not by changes in LAI. When we include the predicted LAI, the overall photosynthesis is higher by only approximately $1 \,\mu$ mol m⁻² s⁻¹ during the dry season. However, if we include a leaf age adjusting factor (Appendix C), the assimilation rate is lower by an average of 1.5 μ mol m⁻² s⁻¹ through-

out the year. The largest difference $(3.37 \,\mu\text{mol}\,\text{m}^{-2}\,\text{s}^{-1})$ occurs in June, when the new leaves start appearing in response to increased sunlight. The assimilation also peaks later than when using a constant LAI, as new leaves reach peak photosynthetic rates only after a certain period of time. While some ground studies report a decrease in assimilation rate during dry periods (Malhi et al., 1998; Miranda et al., 2005) and during severe drought periods (Phillips et al., 2009), the lower assimilation rate at the start of the dry season has been observed in ground studies (Hutyra et al., 2007; Goulden et al., 2004; Graham et al., 2003; Bonal et al., 2008; da Rocha et al., 2004) but previous models were unable to predict this pattern. The hypothesis advanced by one of these studies (Hutyra et al., 2007) was that the emergence of new leaves at the start of the dry season would create this pattern, which is quantitatively supported by the leaf demography predicted by our model.

5 Concluding remarks

We present a simple phenology model for the Amazon basin, which we fitted to 5 years of MODIS LAI data. We showed that our model reproduced the observed increase in LAI during the dry season as a response to an increase in direct solar radiation. Our model parameters provided further information about the vegetation in the Amazon basin. The



Fig. 7. Parameter uncertainty derived from the posterior distribution, expressed as 95 % confidence intervals relative to posterior means for (a) direct PAR compensation point, C_{direct} ; (b) diffuse PAR compensation point, C_{diffuse} ; (c) delay in vegetation response to changes in PAR, p; (d) maximum number of leaves that can be added over a month, gain_{max}; (e) age after which age related decay starts, age_{crit}; (f) background decay constant, μ_0 and (g) age-related decay constant, μ_1 .

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Fig. 8. Estimated frequency distribution of leaf ages (months) over (a) the eastern Amazon, where leaves are typically short-lived, (b) over the semi-deciduous Amazon, and (c) over the evergreen central Amazon basin. All locations are the same as for Fig. 4.



Fig. 9. Gross carbon assimilation calculated using a simple carbon model for a constant LAI, model LAI and model LAI including the temporal variations in leaf age. The values presented are monthly means over a 5 year period at one location (8° N 55° W). All values are μ mol m⁻¹ s⁻¹.

model also provided leaf demography estimates, which can be used to improve predictions of the seasonal carbon cycle, which we demonstrated in principle using a simple carbon model. We showed that using our predicted leaf demography explains observed decrease in carbon assimilation at the start of the dry season. Carbon fixation is only one of many examples of leaf-age-dependent processes of which our current understanding is hampered by incomplete knowledge of leaf demography. Recent work has shown that including a better description of leaf phenology in Earth system models can significantly revise estimates of land surface warming (Bounoua et al., 2010). The demographic model presented here can be used to predict responses of Amazon leaf demography to future changes in climate and could be extended to include other tropical regions where leaf phenology is driven partially or wholly by soil moisture. We therefore anticipate that the insights afforded by our analysis will have far-reaching implications for improving current understanding of the natural carbon cycle in the Amazon and elsewhere.

Appendix A

Model structure

A1 Attenuation Coefficient

To quantify the attenuation of photosynthetic active radiation (PAR) as it passes through a forest canopy, we used Beer's law:

$$I = I_0 e^{-\alpha L},\tag{A1}$$

where I_0 is the incoming PAR (Wm⁻²) at the top of the canopy, I is the light level at layer L inside the canopy and α is the light attenuation coefficient. We assume a vertically homogeneous canopy with no leaf clumping. The light attenuation coefficient is a function of the solar inclination angle, ϕ . For a random distribution of leaf angles this is equal to $\alpha = \frac{0.5}{\sin \phi}$. We use a homogeneous canopy with no leaf clumping and a random leaf angle distribution as these assumptions are valid for canopies at a large scale. Because we are using daily time scales, we can, for simplicity, calculate the attenuation coefficient at its maximum value for a solar angle equal to 90°. For non-directional (diffuse) radiation the attenuation coefficient can be calculated as the median value over all possible radiation angles. The median is used here rather than the mean to avoid the excessive influence of very small or very large angles.

Location	Measured litterfall	Predicted leaf loss	Reference
2°51′ S 54°58′ W	0.25-0.81(8.16)	0.47-0.62	Malhado et al. (2009), Doughty and Goulden (2008)
0°25′–1°30′ S, 72°30′–70°40′ W	0.24-0.55	0.03-0.11	Lips and Duivenvoorden (1996)
11°24'S 55°19′ W	0.11-1.42	0.002-0.14	Sanches et al. (2008)
4°45′–05°30′ N 60°30′–61°22′ W	0.17-0.72	0.52-0.56	Dezzeo and Chacon (2006)

Table A1. Comparison of leaf litterfall as predicted by the model and ground-based measurements across the Amazon basin.

A2 Soil Water Target

Soil water is widely recognised as a primary constraint in LAI, and the seasonality of LAI, in many different vegetation types around the world. For an individual tree, a greater leaf area implies a higher stomatal conductance and hence a greater potential rate of evapotranspiration. If this potential rate cannot readily be met with available soil moisture, the tree can either keep stomata open, risking excessive water loss, cavitation and hence drought death; or close the stomata, which greatly reduces water loss, but also shuts down photosynthesis and risks leaves overheating, causing permanent damage to the leaves; or the tree can reduce its total leaf area, thus allowing for individual leaves to maintain evapotranspiration and photosynthesis without excessive water use overall (McDowell et al., 2008). The first two options are short term responses to unpredictable drought, whereas losing leaves is the more sensible response to the seasonal variation in soil moisture that drives leaf phenology. Therefore, we assume that where the current LAI exceeds the waterlimited LAI $\ensuremath{\mathsf{LAI}}\xspace^{\ensuremath{\mathsf{RAI}}\xspace}\xspace$, trees would drop leaves in order to reach LAI_{targ}^{water}, starting with the oldest leaves, to produce a leaf demography that is sustainable in the long term (see main text).

Based on these assumptions, we specify a relationship between soil moisture and LAI^{water} as:

$$LAI_{targ}^{water} = \beta_1 + \beta_2 W_s, \tag{A2}$$

where W_s is the percentage soil moisture, and β_1 and β_2 are empirical coefficients. If the current LAI is above this target, trees will drop leaves until they reach LAI^{water}.

We assume trees actively drop leaves when LAI exceeds LAI_{targ}^{water} , but do not do so when LAI exceeds LAI_{targ}^{light} . Where LAI exceeds LAI_{targ}^{light} , the leaves receiving the least light will be below their compensation point and so be a net sink, rather than source, of carbon. However, the magnitude of this sink will be relatively small, when compared with the magnitude of the source from the leaves receiving more light. Also, trees can store substantial amounts of labile carbon, which can be used to offset short term deficits in carbon fixation. Therefore, where LAI exceeds LAI_{targ}^{light} , the excess leaves are unlikely to be a cause of whole plant stress or death. In contrast, as outlined above, where LAI exceeds LAI_{targ}^{light} the lack of

water affects the entire tree, and so could cause damage to all leaves, or to the whole plant through cavitation. In addition, very few trees store an amount of water that is significant in comparison to daily water use. Therefore, when LAI exceeds LAI^{water}, the excess leaves are an immediate source of danger to the whole tree.

We fit the model using the method as described in Sect. B below, using NCEP/NCAR reanalysis derived soil moisture data (Kalnay et al., 1996). The values obtained for the β_1 and β_2 parameters result in a water limited target LAI^{water}_{targ} that is higher than the light limited target LAI^{light}_{targ} at all locations throughout the year, This implies that vegetation is never water limited. As a result we set LAI_{targ} = LAI^{water}_{targ}, which results in the model structure described in Fig. 2.

Appendix B

Parameter estimation

The model has 7 free parameters (Table 1), which we fit independently for each location, resulting in 840 parameters.

Our aim was to estimate model parameters for location **x**, which we denote as the vector $\theta_{\mathbf{x}}$, given the MODIS data for location **x**, which we denote $Z_{\mathbf{x}}$. To do this we used a Bayesian approach, seeking to estimate the posterior probability distribution of $\theta_{\mathbf{x}}$, given $Z_{\mathbf{x}}$, which we denote $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})$. The posterior $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})$ is proportional to the product of the likelihood $L(Z_{\mathbf{x}}|\theta_{\mathbf{x}})$, and the prior $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})$ such that $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}}) \propto L(Z_{\mathbf{x}}|\theta_{\mathbf{x}}) p(\theta_{\mathbf{x}})$. Often, $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})$ covers too large a region of parameter space to be evaluated completely. In this case we instead use sampling methods, which provide a set of random samples of $\theta_{\mathbf{x}}$ drawn from $p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})$. We used a Metropolis-Hastings Markov Chain Monte Carlo (MCMC) sampling routine (Gilks, 1996.).

We define the likelihood as

$$l(Z_{\mathbf{x}}|\theta_{\mathbf{x}}) = \sum_{t(\mathbf{x})} \ln[n(\mathrm{LAI}_{\mathrm{obs}}(\mathbf{x}, t), \mathrm{LAI}_{\mathrm{pred}}(\mathbf{x}, t, \theta_{\mathbf{x}}), \sigma_{\mathbf{x}})] \qquad (B1)$$

where $l(Z_{\mathbf{x}}|\theta_{\mathbf{x}})$ is the log-likelihood; $LAI_{pred}(\mathbf{x}, t, \theta_{\mathbf{x}})$ is the predicted LAI at location \mathbf{x} at time t (this depends on the model parameters $\theta_{\mathbf{x}}$); $LAI_{obs}(\mathbf{x}, t)$ is the observed MODIS LAI at location \mathbf{x} at time t; and $n(LAI_{obs}(\mathbf{x}, t),$ LAI_{pred}($\mathbf{x}, t, \theta_{\mathbf{x}}$), $\sigma_{\mathbf{x}}$) denotes the probability density for observing LAI_{obs}(\mathbf{x}, t) given a normal distribution with mean LAI_{pred}($\mathbf{x}, t, \theta_{\mathbf{x}}$) and standard deviation $\sigma_{\mathbf{x}}$ where $\sigma_{\mathbf{x}}$ is a parameter that specifies the magnitude of unexplained variation in LAI. Eq. B1 represents a loop over all times *t* for which observed LAI were available for location \mathbf{x} (this set of times is denoted *t*(\mathbf{x}) in Eq. B1).

We initially used non-informative, uniform priors for all parameters. We assumed that, a priori, all parameter combinations were equally likely. However, we found that, with non-informative priors on all parameters, we could not properly constrain all parameters for all locations. Therefore, we assigned an informative prior on one of the parameters affecting the rate of gain and loss of leaves. We did this for one parameter only to keep the overall influence of priors to a minimum. We base our prior on an extensive study (Reich et al., 2004) of leaf lifespan (time after which all leaves are dead) in the northern Amazon. To define the prior we use the mean (2.35 years) and standard deviation (0.18 years) of these measurements. In our model, we define leaf lifespan $\tau_{95,\mathbf{x}}$ as the leaf age at which only 5 % of leaves remain alive. This is a function of three model parameters: $\mu_{0,\mathbf{x}}$, $\mu_{1,\mathbf{x}}$ and $a_{\text{crit},\mathbf{x}}$. Using this prior, the posterior in the MCMC sampling becomes:

$$\ln(p(\theta_{\mathbf{x}}|Z_{\mathbf{x}})) = l(Z_{\mathbf{x}}|\theta_{\mathbf{x}}) + \ln(n(\tau_{95,\mathbf{x}},\hat{\tau},\sigma_{\tau})),$$
(B2)

where $n(\tau_{95,\mathbf{x}}, \hat{\tau}, \sigma_{\tau})$ is the probability density for $\theta_{95,\mathbf{x}}$ assuming that it is drawn from a normal distribution with mean $\hat{\tau}$ and standard deviation σ_{τ} . This measure is only proportional to the posterior because it does not take into account an integration constant. However, when using MCMC sampling this is unimportant because the constant cancels out when calculating the probability for acceptance and rejection. With this simple leaf lifespan constraint in place, we found that all parameters converged for all locations and that leaf lifespan varies substantially across the region, showing that the parameter estimates were not overly influenced by the prior (otherwise all leaf lifespans would have converged on $\hat{\tau}$). Also, parameters not directly affecting $\tau_{95,\mathbf{x}}$ were estimated to have reasonable values, implying that the model structure, MODIS data, the prior on leaf lifespan, and previous knowledge on other parameters, were all consistent with each other.

We used Metropolis-Hastings MCMC sampling (hereafter MH-MCMC) to provide a set of 600 random samples from $p(\theta_x | Z_x)$. MH-MCMC is a widely recognised and simple, albeit computational heavy, method to provide samples from the posterior parameter distribution. MH-MCMC samples this distribution simply by proposing new parameter sets and accepting or rejecting these on the basis of their posteriors, according to a standard Metropolis criterion. Given a sufficient number of steps, the random walk reaches a quasiequilibrium, after which the average properties of the walk (e.g. the mean and standard deviation as measured against

any one parameter) no longer change. After this quasiequilibrium has been reached, the current position of the algorithm at any one time constitutes a random sample from the posterior (Gelman, 2004). However, there is a great deal of freedom in exactly how to carry out MH-MCMC for a particular case, especially in how new parameter sets are proposed. For our analysis we used "Filzbach" ¹, a previously tested and robust algorithm.

We allowed 60 000 iterations for the burn-in (the period to quasi-equilibrium), then sampled every 100 iterations from a further 60 000 total iterations, thus providing our 600 samples from $p(\theta_x|Z_x)$. From these samples we calculated, for each parameter, the posterior mean, and 95 confidence intervals. All values reported in the paper are posterior means. We have used parameter sets drawn from this distribution to obtain LAI values (Fig. 2) and then calculate average LAI values and upper and lower 95% confidence bounds for all values.

B1 Generating predicted LAI values

In order to both parameterize, and run simulations of, our model, it was necessary to generate predicted LAI values for each location \mathbf{x} and time t, (e.g. see the likelihood defined above). Importantly, as a demographic model, our model as defined in the main text produces rates of change of the LAI held within different leaf age classes. Therefore, to produce LAI($\mathbf{x}, t, \theta_{\mathbf{x}}$) it is necessary to first, set an initial leaf age distribution, and second, simulate the model forward in time from this initial state, until we reach time t. The schematic for this is given in Fig. 2. To make our predictions for a given parameter set θ_x , we set the initial leaf age distribution as LAI $(a, \mathbf{x}, t, \theta_{\mathbf{x}}) = 0$ for all *a* (corresponding to no leaves) then simulated the year 2000 10 times over, each time beginning the simulated year, using as the initial state the leaf age distribution from the end of the previous simulated year. This acted to spin up the model to produce a reasonable initial leaf age distribution consistent with current parameters. After the spin-up, we then simulated the model forward in time in steps of 8 days, keeping note of LAI($\mathbf{x}, t, \theta_{\mathbf{x}}$) for all t where we had observations of LAI with which to compare the predictions. Note that, although for the parameter estimation we only used the total LAI, $LAI(x, t, \theta_x)$, the model can only be simulated by keeping track of the LAI within each age class, LAI $(a, \mathbf{x}, t, \theta_{\mathbf{x}})$. Thus predicted leaf age distributions emerge as a natural outcome of applying our model.

¹Filzbach, a software library for carrying out Metropolis-Hastings Markov chain Monte Carlo parameter estimation in C++ or C#. Filzbach is under development in the Computational Science lab at Microsoft Research Cambridge and is available for download, complete with a suite of example uses, via http://research.microsoft. com/en-us/groups/ecology/ecotechandtools.aspx

Appendix C

Carbon Assimilation Model

To illustrate the impact that our model results have on the carbon cycle in the Amazon, we use a simple canopy model to describe leaf photosynthesis rates. We assume that the only limitation to photosynthesis is light availability, so that carbon assimilation rates are a linear function of incoming PAR:

$$A(I) = \begin{cases} \phi I - q , I < I_{\max} \\ A_{\max} , I > I_{\max} \end{cases}$$
(C1)

where A_{max} is the maximum assimilation rate that occurs after photosynthesis reaches saturation with light for PAR levels above I_{max} . We use literature values (Kubiske and Pregitzer, 1996; Riddoch et al., 1991; Langenheim et al., 1984; Miranda et al., 2005; Hutyra et al., 2007; Kitajima et al., 1997) to obtain an A_{max} of 6.12 µmol m⁻² s⁻¹ and I_{max} equal to 150 µmol m⁻² s⁻¹ (Hutyra et al., 2007). The values cited above have been measured for various locations, species and light environments which we have averaged to obtain a canopy scale value. We then use these values to calculate ϕ and q, by assuming that assimilation is zero at a light level equal to the compensation point C_{direct} .

We use our posterior age distributions to correct photosynthetic rates for the effects of leaf ageing. Studies have shown (Kitajima et al., 1997; Doughty and Goulden, 2008) that in tropical systems photosynthesis rates peak a few weeks after budburst and that measured rates decline with age, reaching half the peak value for leaves older than 1 year. Of course, these figures do not reflect the large variation in leaf lifespan in the Amazon. It has been observed that longer lived leaves show a slower decline in assimilation rates with age and are also slower to reach peak rates (Kitajima et al., 1997).

To account for these changes we use an age correction factor. Assuming that the values defined above are correct for mature leaves, then a population composed entirely of mature foliage will have an age factor of 1, while populations with a combination of young, mature and old leaves will have a factor less than 1. We define this factor as

$$\gamma_{\text{age}} = f_{\text{new}} A_{\text{new}} + f_{\text{mat}} A_{\text{mat}} + f_{\text{old}} A_{\text{old}}, \qquad (C2)$$

where f_{new} , f_{mat} and f_{old} are the fractions of young (age < 0.07 τ_{95}), mature and old (age>age_{min}) leaves respectively. The corresponding adjusting factors are equal to 0.05 for young leaves, 1 for mature and 0.5 for old. We assume that the A_{max} and I_{max} values used above are valid for fully mature leaves.

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Appendix D

Predicted leaf litterfall

We compare model predictions against ground-based measurements of leaf litterfall 9Table A1. This provides an evaluation of the model parameters that is independent of the LAI data used to parameterise the model. All litterfall measurements have been converted from mass units $(gm^{-2} month^{-1})$ to area units $(m^2 m^{-2} month^{-1})$ using either the leaf mass per unit area value given by (Fyllas et al., 2009) (94.85 gm⁻²) or the specific value for that site if any is given in the study (mentioned in brackets).

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