

Hoeffding's inequality in game-theoretic probability

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Abstract

This note makes the obvious observation that Hoeffding's original proof of his inequality remains valid in the game-theoretic framework. All details are spelled out for the convenience of future reference.

1 Introduction

The game-theoretic approach to probability was started by von Mises and greatly advanced by Ville [5]; however, it has been overshadowed by Kolmogorov's measure-theoretic approach [3]. The relatively recent book [4] contains game-theoretic versions of several results of probability theory, and it argues that the game-theoretic versions have important advantages over the conventional measure-theoretic versions. However, [4] does not contain any large-deviation inequalities. This note fills the gap by stating the game-theoretic version of Hoeffding's inequality ([2], Theorem 2).

2 Hoeffding's supermartingale

This section presents perhaps the most useful product of Hoeffding's method, a non-negative supermartingale starting from 1. This supermartingale will easily yield Hoeffding's inequality in the following section.

This is a version of the basic forecasting protocol from [4]:

GAME OF FORECASTING BOUNDED VARIABLES

Players: Sceptic, Forecaster, Reality

Protocol:

Sceptic announces $\mathcal{K}_0 \in \mathbb{R}$.

FOR $n = 1, 2, \dots$:

Forecaster announces interval $[a_n, b_n] \subseteq \mathbb{R}$ and number $\mu_n \in (a, b)$.

Sceptic announces $M_n \in \mathbb{R}$.
Reality announces $x_n \in [a_n, b_n]$.
Sceptic announces $\mathcal{K}_n \leq \mathcal{K}_{n-1} + M_n(x_n - \mu_n)$.

On each round n of the game Forecaster outputs an interval $[a_n, b_n]$ which, in his opinion, will cover the actual observation x_n to be chosen by Reality, and also outputs his expectation μ_n for x_n . The forecasts are being tested by Sceptic, who is allowed to gamble against them. The expectation μ_n is interpreted as the price of a ticket which pays x_n after Reality's move becomes known; Sceptic is allowed to buy any number M_n , positive, zero, or negative, of such tickets. When x_n falls outside $[a_n, b_n]$, Sceptic becomes infinitely rich; without loss of generality we include the requirement $x_n \in [a_n, b_n]$ in the protocol; furthermore, we will always assume that $\mu_n \in (a_n, b_n)$. Sceptic is allowed to choose his initial capital \mathcal{K}_0 and is allowed to throw away part of his money at the end of each round.

It is important that the game of forecasting bounded variables is a perfect-information game: each player can see the other players' moves before making his or her (Forecaster and Sceptic are male and Reality is female) own move; there is no randomness in the protocol.

A *process* is a real-valued function defined on all finite sequences $(a_1, b_1, \mu_1, x_1, \dots, a_N, b_N, \mu_N, x_N)$, $N = 0, 1, \dots$, of Forecaster's and Reality's moves in the game of forecasting bounded variables. If we fix a strategy for Sceptic, Sceptic's capital \mathcal{K}_N , $N = 0, 1, \dots$, become a function of Forecaster's and Reality's previous moves; in other words, Sceptic's capital becomes a process. The processes that can be obtained this way are called (game-theoretic) *supermartingales*.

The following theorem is essentially inequality (4.16) in [2].

Theorem 1 *For any $h \in \mathbb{R}$, the process*

$$\prod_{n=1}^N \exp \left(h(x_n - \mu_n) - \frac{h^2}{8}(b_n - a_n)^2 \right)$$

is a supermartingale.

Proof Assume, without loss of generality, that Forecaster is additionally required to always set $\mu_n := 0$. (Adding the same constant to a_n , b_n , and μ_n will not change anything for Sceptic.) Now we have $a_n < 0 < b_n$.

It suffices to prove that on round n Sceptic can make a capital of \mathcal{K} into a capital of at least

$$\mathcal{K} \exp \left(hx_n - \frac{h^2}{8}(b_n - a_n)^2 \right);$$

in other words, that he can obtain a payoff of at least

$$\exp \left(hx_n - \frac{h^2}{8}(b_n - a_n)^2 \right) - 1$$

using the available tickets (paying x_n and costing 0). This will follow from the inequality

$$\exp\left(hx_n - \frac{h^2}{8}(b_n - a_n)^2\right) - 1 \leq x_n \frac{e^{hb_n} - e^{ha_n}}{b_n - a_n} \exp\left(-\frac{h^2}{8}(b_n - a_n)^2\right),$$

which can be rewritten as

$$\exp(hx_n) \leq \exp\left(\frac{h^2}{8}(b_n - a_n)^2\right) + x_n \frac{e^{hb_n} - e^{ha_n}}{b_n - a_n}. \quad (1)$$

Our goal is to prove (1). By the convexity of the function \exp , it suffices to prove

$$\frac{x_n - a_n}{b_n - a_n} e^{hb_n} + \frac{b_n - x_n}{b_n - a_n} e^{ha_n} \leq \exp\left(\frac{h^2}{8}(b_n - a_n)^2\right) + x_n \frac{e^{hb_n} - e^{ha_n}}{b_n - a_n},$$

i.e.,

$$\frac{b_n e^{ha_n} - a_n e^{hb_n}}{b_n - a_n} \leq \exp\left(\frac{h^2}{8}(b_n - a_n)^2\right),$$

i.e.,

$$\ln(b_n e^{ha_n} - a_n e^{hb_n}) \leq \frac{h^2}{8}(b_n - a_n)^2 + \ln(b_n - a_n). \quad (2)$$

The derivative of the left-hand side of (2) is

$$\frac{a_n b_n e^{ha_n} - a_n b_n e^{hb_n}}{b_n e^{ha_n} - a_n e^{hb_n}}$$

and the second derivative, after cancellations and regrouping, is

$$(b_n - a_n)^2 \frac{(b_n e^{ha_n})(-a_n e^{hb_n})}{(b_n e^{ha_n} - a_n e^{hb_n})^2}.$$

The last ratio is of the form $u(1-u)$ where $0 < u < 1$. Hence it does not exceed $1/4$, and the second derivative itself does not exceed $(b_n - a_n)^2/4$. Inequality (2) now follows from the second-order Taylor expansion of the left-hand side around $h = 0$. \blacksquare

The relation between the game-theoretic and measure-theoretic approaches to probability is described in [4], Chapter 8. Intuitively, the generality of the game-theoretic protocol stems from the fact that Forecaster is not asked to produce a full-blown probability forecast for x_n : only the elements (a_n, b_n, μ_n) that we really need for our mathematical result enter the game of forecasting bounded variables. Besides, the players are allowed to react to each other moves; in particular, Reality may react to Forecaster's moves and both Reality and Forecaster may react to Sceptic's moves (the latter is important in applications to defensive forecasting: see, e.g., [6]). It is remarkable that many measure-theoretic proofs carry over in a straightforward manner to game-theoretic probability.

3 Hoeffding's inequality

We start from the definition of upper probability, a game-theoretic counterpart (along with lower probability) of the standard measure-theoretic notion of probability. Suppose the game of forecasting bounded variables lasts a known number N of rounds. (See [4] for the general definition.) The *sample space* is the set of all sequences $(a_1, b_1, \mu_1, x_1, \dots, a_N, b_N, \mu_N, x_N)$ of Forecaster's and Reality's moves in the game. An *event* is a subset of the sample space. The *upper probability* of an event E is the infimum of the initial value of non-negative supermartingales that take value at least 1 on E . (See [4], Chapter 8, for a demonstration that this definition agrees with measure-theoretic probability.)

Theorem 1 immediately gives Hoeffding's inequality (cf. [2], the proof of Theorem 2) when combined with the definition of game-theoretic probability:

Corollary 1 *Suppose the game of forecasting bounded variables lasts a fixed number N of rounds. If all a_n and b_n are given in advance and $t > 0$ is a known constant, the upper probability of the event*

$$\frac{1}{N} \sum_{n=1}^N (x_n - \mu_n) \geq t \tag{3}$$

does not exceed

$$e^{-2N^2 t^2 / C},$$

where $C := \sum_{n=1}^N (b_n - a_n)^2$.

(The reader will see that it is sufficient for Sceptic to know only C at the start of the game, not the individual a_n and b_n .)

Proof The supermartingale of Theorem 1 starts from 1 and achieves

$$\prod_{n=1}^N \exp \left(h(x_n - \mu_n) - \frac{h^2}{8} (b_n - a_n)^2 \right) \geq \exp \left(hNt - \frac{h^2}{8} C \right) \tag{4}$$

on the event (3). The right-hand side of (4) attains its maximum at $h := 4Nt/C$, which gives the statement of the corollary. ■

Remark The measure-theoretic counterpart of Corollary 1 is sometimes referred to as the Hoeffding–Azuma inequality, in honour of Kazuoki Azuma [1]. The martingale version, however, is also stated in Hoeffding's paper ([2], the end of Section 2).

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