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1 Variance in biomass-allocation fractions is explained by distribution 2 in European trees

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- 16 Summary
- 17 Intraspecific variability in ecological traits confers the ability of a species to adapt to an • ever-changing environment. Fractions of biomass allocation in plants (BAFs) represent 18 19 both ecological traits and direct expressions of investment strategies and so have 20 important implications on plant fitness particularly under current global change. 21 • We combined data on BAFs of trees in more than 10000 forest plots with their 22 distributions in Europe. We aimed to test whether plant species with wider distributions 23 have more or less variable intraspecific variance of BAFs foliage-woody biomass and 24 shoot-root ratios than species with limited distribution. 25 Irrespective of corrections for tree age and phylogenetic relatedness, the standard • deviation in BAFs was up to three times higher in species with most extensive 26 27 distributions than in those with least extensive distribution due to a higher genetic 28 diversity. Variance in BAFs also increased with latitude. 29 We show that a combination of 36% tree genetic diversity and 64% environmental • variability explains variance in BAFs and imply that changes in genetic diversity occur 30 31 quickly. Genetic diversity should thus play a key role in regulating species responses to 32 future climate change. Loss of habitat, even if transient, could induce a loss of genetic 33 diversity and hinder species survival. 34 Key-words: allometry; effective population size; extinction; genetic diversity; plant allometry; 35 36 plant morphology; woody plants
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39 Introduction

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41 The strategies of fitness and growth of sessile organisms are largely determined by biomass 42 allocation (Hodge, 2004; Poorter et al., 2012; Veresoglou et al., 2017), which in turn determines 43 the long-term morphology of the individual. The morphology of sessile organisms has been 44 studied the most in plants (Kokko, 2007). Plant morphology can be summarised in several ways, 45 each with distinct strengths and limitations. Poorter and Sack (2012) reviewed these methods and 46 concluded that fractions of biomass allocation, especially after correcting for confounding 47 allocation parameters such as size, represent a particularly effective measure. Plant morphology 48 has high functional importance, so the influence of environment on morphology has been well 49 studied. Some representative parameters that have been studied extensively for their influence on 50 plant allometry are the size of individual plants (Reich et al., 2005; Poorter et al., 2012), 51 historical environmental abiotic parameters such as stressors (e.g. uniform stress hypothesis) 52 (Mogan & Gannell, 1994; Dean et al., 2002), shading (Lusk et al., 2008; Duursma et al., 2010; 53 Forster et al., 2011), temperature (Reich et al., 2014), precipitation (McCarthy & Enquist, 2007), 54 biotic interactions such as competition and diversity (Forrester et al., 2017a); the growth form of 55 plants (Wyka et al., 2013) and species-specific ontogeny (Forrester et al., 2017b). Plant age, 56 however, is an influential factor that is often not available for integration into such analyses 57 (Nikklas, 1997; Nikklas et al., 2003; Barthélémy & Caraglio, 2007; Duursma et al., 2010). 58 Analysts often use plant size as a proxy of age (Nikklas, 2004; Bowman et al., 2013). 59 Biomass-allocation fractions (BAFs) describe biomass ratios, usually at a logarithmic 60 scale, of plant organs. BAFs are measureable to an individual level and thus constitute plant 61 traits (Müller et al., 2000; Reich et al., 2003). Trait variance is often comparable to or even more 62 important than trait means (Messier et al., 2010; Bolnick et al., 2011; Violle et al., 2012). High 63 variance in trait values could facilitate, for example, the evolution and adaptation of a species to 64 new environmental settings (Bolnick et al., 2011). Quantifying variation in BAFs could improve 65 modelling uncertainty in the standing biomass of woody habitats because most of our estimates 66 of standing biomass are projections of allometric equations (Chave et al., 2005; Muukkonen, 2007). 67

68 Plant morphology has been described as an equilibrium between constraints to plant
69 growth and exogenous environmental stressors (Barthélémy & Caraglio, 2007). Constraints to

70 plant growth and other intrinsic factors can negate the influence of the environment to varying 71 degrees, i.e. environmental conditions tend to shift BAFs against the stabilising influence of 72 intrinsic factors. We would then expect that BAFs would be more variable in plant species that 73 experience extreme environmental conditions more often. Rever et al. (2013) argued that 74 extreme conditions occur mainly at the edges of the distribution of a species. Source-sink 75 population dynamics describe instances where species only occur in an area because of a 76 constant influx of propagules from surrounding areas where the species grows better. At the 77 edges of the distribution of a species it is more likely to observe source-sink population 78 dynamics (Remeš, 2000) than in the kernel of the distribution which could induce unique 79 phenotypes. Plant species that have a small distribution should more frequently experience such 80 "extreme edge" conditions so plants with smaller distributions may have the most variable BAFs. 81 The reason is that environmental heterogeneity mainly increases the variance in BAFs in these 82 species (Fig. 1a). We thus hypothesize that plant species with smaller distributions have more 83 variable BAFs (hypothesis one). Alternatively, the higher effective population sizes and genetic 84 variability could allow plant species that have extensive distributions to be those showing the 85 highest intraspecific trait variance, including BAFs. The reason could be that environmental 86 variability acts independently of distribution but more populous species show more diverse 87 BAFs due to a higher genetic diversity (Fig. 1b). As a result, plant species with smaller 88 distributions could vary less in their BAFs (hypothesis two). A final expectation is that variance 89 in BAFs should indicate the ability of species to adapt to local environmental conditions, so a 90 high variance suggests rapid evolution. Stapley et al. (2017) reported considerably lower 91 recombination rates in gymnosperms than angiosperms. The variance in BAFs should thus be 92 lower in conifers, compared to angiosperms (hypothesis three).

93 Most of our existing understanding on the way traits of woody species vary with genetic 94 diversity and the environment originates from provenance tests, where plants differing in their 95 origin are grown under common environmental conditions (Thompson et al. 2008; Wang et al. 96 2010). Some of the limitations of provenance test relate to the choice of the common 97 environment (Leites et al. 2012), the time they require for long-lived species to grow and thus the 98 logistics of destructively harvesting them (e.g. to assess BAFs). Here we worked on an 99 observational approach which synthesizes across existing BAF measurements from records of 100 destructive tree harvests in Europe (Fig. 1). Schepaschenko et al. (2017) have recently released

101 two large data sets detailing the biomass fractions of many trees that had been destructively

102 harvested between 1930 and 2014 in Eurasia. We combined this data set with information on the

103 distributions of many of these plant species in Europe from Mauri et al. (Mauri *et al.*, 2017) to

address the above mentioned three hypotheses.

105

106 Materials and Methods

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108 Sources of data

109 We used the two data sets published by Schepaschenko et al. (2017) for our main analysis. The 110 Biomass-tree data set provided information on biomass fractions following destructive 111 harvesting for 9613 trees, mainly in Europe. The Biomass-plot data set provided information on 112 biomass fractions for 10351 plots distributed throughout Eurasia, each of which provided 113 information for cumulative biomass of two or more trees. The two data sets synthesized across approximately 1200 experiments over the period 1930-2014 and contained information on 114 115 location and age of the trees as well as biomass information on several different fractions. We 116 used biomass information on foliage vs. woody aboveground biomass for our main analyses. We 117 also used root:shoot ratios as part of our sensitivity analyses. We extracted distributions of tree 118 species in Europe from Mauri et al. (Mauri et al., 2017). Mauri et al. (2017) only described tree 119 distributions in Europe, so we limited our analysis to trees that occur mainly in Europe. We used 120 QGis v 2.12.3 to estimate the size of the polygonal envelopes provided by Mauri et al. (2017). 121 We used Phylocom v 4.2 to reconstruct the phylogenetic relationships of the tree species in our 122 analysis.

123 We worked with an aggregate of 80 species in our analysis of European trees. A list of 124 the species and the phylogenetic reconstruction used to correct for phylogenetic relatedness can 125 be found in Table S1. The three data sets varied considerably in terms of their resolution and 126 their suitability for the different analyses. The cumulative analysis of the *Biomass-tree* data set 127 used biomass and age information for 4719 trees from 42 species (median observations per species, 10.5; 1st quartile, 7; 3rd quartile, 78.5). The analysis of the *Biomass-plot* data set used 128 129 information for 3898 plot entries describing 63 species (median observations of plot per species, 15; 1st quartile, 8.5; 3rd quartile, 46) making it particularly suitable for assessing how BAFs are 130 131 influenced by environmental heterogeneity. The analysis of individual stands in the Biomass-tree 132 data set used an aggregate of 1854 tree observations from 40 species (median observations per

- 133 species, 8.5; 1st quartile, 7; 3rd quartile, 33.25). Because for the analysis on individual stands (i.e.
- 134 plots) in the *Biomass-tree* data set we only used per species information of trees in a single stand
- 135 (i.e. plot with most tree observations in the *Biomass-tree* data set), this made it particularly
- suitable for quantifying the influence upon BAFs of genetic diversity. The *Biomass-tree* data set
- 137 contained proportionally more observations located in Europe than *Biomass-plot* data set did
- 138 (Fig.2 inserts).
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140 Rationale for the analyses

141 We modelled three sources of variance in BAFs, age, environment and genetic diversity. We 142 inferred genetic diversity from genetic variability (i.e. the degree to which the genetic 143 characteristics of populations vary) using a phenomenological approach that determined the 144 variance of morphological characteristics in populations after controlling for sources of variance 145 that were not of genetic origin such as age, environment and latitude. To control for age-related 146 differences in BAFs we used the slopes of linear models with BAFs as response variables and 147 age as predictor. To assess the fraction of BAFs that is explained by genetic diversity we 148 assessed how variance in BAFs scaled with tree species distribution in tree individuals which 149 belonged to the same species and stand after correcting them for tree age (individual stands in the 150 Biomass-tree data set; Fig. 1b). To assess the fraction of BAFs that is explained by 151 environmental variability we assumed that environmental variability is mainly due to 152 latitudinal range (de Frenne et al. 2013; we provide more details at the section poling effect 153 sizes) and quantified how much the variance in BAFs changed with each degree of latitudinal 154 range (i.e. latitudinal breadth - to avoid an overrepresentation of stands we used the *Biomass*-155 plot dataset).

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157 Statistical analyses

Our analyses only considered tree species (which were the unit of our analysis) for which a minimum of five observations of age, foliage biomass and aboveground woody biomass were available, because we needed a minimum of three observations to fit a linear model and then two additional residuals to yield meaningful estimates of variance for the fit of the regression line (results from a sensitivity test on the inclusion threshold are presented in Appendix 1 and Fig.

- 163 S1). To extract the fraction of variance in BAFs that was not due to age differences we fitted
- 164 linear models with BAFs as a response variable and age as predictor and then quantified the
- 165 standard deviation of the residuals. We explain the procedure in higher detail below (also

166 Appendix 2; Table S2):

167 1. We calculated the natural log response ratio (logRR) of foliage over aboveground woody

168 biomass $\left(\log RR = \ln \left(\frac{m_{foliage}}{m_{woody}}\right)$, where *m* stands for biomass) for each individual tree in the case of

- 169 the *Biomass-tree* data base or plot in the case of the *Biomass-plot* data base (Appendix 2). In the
- 170 form of a sensitivity analysis we also analysed logRR of root over shoot BAFs. LogRR
- 171 represents a widely used nonstandard effect size in synthesis studies. A large logRR indicates a
- 172 higher investment in foliage than woody biomass.

173 2. We fitted a linear model with logRR as a response variable and age as the sole predictor. The

- 174 single most important idiosyncratic cause of variability in BAFs is age (Nikklas *et al.*, 2003;
- 175 Nikklas, 2004; Barthélémy & Caraglio, 2007; Duursma et al., 2010). To correct our data for this
- 176 source of variance, we extracted the residuals of the linear model and assessed their standard
- 177 deviations. We assumed a first-order linear correlation between logRR and age using the
- standard deviation of the residuals as a measure of the variance, in agreement with preliminaryanalyses (Appendix 2).
- 3. We correlated our metric of variance of the relationship between BAFs and age with thedistribution of the plants.

182 In the first of our three analyses (Table S2) we used all observations of individual trees in 183 the *Biomass-tree* data set, which consisted of multiple trees per stand and multiple stands per 184 species. We then used information for all plots in the *Biomass-plot* data set, consisting of a single 185 plot per stand, which allowed us to correct for spatial autocorrelation. Finally, we reanalysed the 186 *Biomass-tree* data set but only using the information for each species in a single stand, with the 187 sole criterion that the stand had provided most observations for that species. This analysis 188 addressed the concern that the cause of higher variance in BAFs was due to a larger distance 189 across observations and thus a higher variability in environmental conditions. We conducted this 190 analysis both with and without phylogenetic correction. We used an analysis of phylogenetically 191 independent contrasts to correct for phylogeny.

192 193 We further compared our estimates of variance between angiosperms and gymnosperms. We thus used *t*-tests, assuming unequal variance.

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195 Sensitivity analyses

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We tested the specificity of our observations to logRR between foliage and aboveground woody biomass by replicating the analysis for logRR between total root biomass vs. total aboveground biomass. The two data sets contained unequal numbers of observations of total root biomass, and the number of species which we could analyse was low. We directly assessed how the variance in biomass fractions scaled with distribution (i.e. independent of age) by directly assessing the standard deviation of biomass-allocation ranges and correlating it with distribution (Appendix 2).

203 We further replicated the analysis with the Legacy Tree Data data set (Radtke *et al.*, 204 2016). To estimate the distribution of North American trees we used information from the United 205 States Department of Agriculture (USDA 2017). This data set did not contain any information on 206 tree age. We used data from the Legacy Tree table and used the variability of logRR foliage over 207 aboveground woody biomass for single locations. Sufficient information was available for only 208 three species, which were assessed for four locations. All three species belonged to the Pinaceae 209 family, had nested distributions and occurred in southeastern North America. We assessed 210 distribution by extracting information from the USDA for the number of states in which the 211 species occurred and regressed this metric against the variance of logRR. The amount of 212 information we extracted from the Legacy Tree Data base and the resolution, given that the size 213 of US states varies considerably, were much lower than in our main analysis so we only used as 214 a means of supporting the main results with a different dataset. Because we did not have accurate 215 distribution range data, we used here non-parametric statistics.

To assess the degree to which our observations were sensitive to the inclusion of invasive tree species we classified tree species in native vs invasive as in Veresoglou et al. (2018; Table S3). We repeated the analysis separately for the subsets of native and invasive species.

219

220 The influence of latitude

Latitude is an influential predictor of variance in plant traits (Heibo et al., 2005; Aerts et al.,

222 2012), and as such it was important for us to show that it was not driving our results. We

assessed the degree to which our patterns were explained by latitude using two analyses. (a) We
first extracted the average latitudes of the distributions of the woody species in Europe reported
by Mauri et al. (2017). We compared these values with the variances in BAFs corrected for plant
age (i.e. the standard deviation of the residuals of the linear model described above). (b) We then

divided Europe into northern and southern Europe. We thus used the average of the two

228 latitudinal extremes of Europe: 82°N for the northern region and 35°N for the southern region,

averaging 58.5°N. We used the subset of points west of 69°W. We separately calculated the

230 variance of BAFs corrected for age for the woody species for which the *Biomass-tree* data set

contained a minimum of five observations both north and south of the latitudinal average

threshold, which we then combined in a new logRR (i.e. northern variance over southern

variance). We calculated the variance for these two sites and used them to recalculate logRR.

234

235 Pooling effect sizes to quantify genetic and environmental variance in BAFs

We partitioned here variance in BAFs into a fraction due to genetic variability and a fraction due
to environmental variability. We used a phenomenological approach for this (Fig. 1; Appendix
4). We made the following three assumptions:

1. The key factor contributing to environmental variability was latitudinal range of the

240 distribution of an organism. In support to this simplistic statement, see de Frenne et al. (2013)

241 demonstrating that some key environmental parameters such as temperature, precipitation, soil

242 pH and human influence covary with latitude. We assumed that there is a first order linear

243 relationship between latitudinal range and environmental variability in BAFs.

244 2. Genetic variability in BAFs increased with distribution at a slope equivalent to that in Fig. 2b.

245 3. The relative importance of the two fractions was assessed for an imaginary species with an

246 "average" latitudinal range which in this case was 19.4° latitude and a mean distribution range of
247 10,000 km².

248 We use the additive principle of variances in statistics to add the two fractions on the assumption

that the observed variance was independent of our sample sizes. Our variance partitioning was

250 carried out for an idealized species having an average latitudinal range of $58.9^{\circ} - 39.5^{\circ} = 19.4^{\circ}$

251 latitudinal range and a distribution of 10.000 km² which represented averages in our dataset and

were also consistent with (Tkach *et al.*, 2008). We estimated the overall form of the model and

we calculated relative effect sizes for mean parameters of distribution and latitude in Europe. Weprovide more information on the analysis in Appendix 4.

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- 259 Results
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The main results are on records throughout Eurasia but these were supported by a 261 262 reanalysis exclusively to the records from Europe. Our main analysis gave consistent results for 263 all three data sets (Fig. 2), namely that variance in BAFs increases with the distribution size of a 264 species. LogRR of foliage vs aboveground woody biomass was significantly positively correlated with distribution in all three tests (Fig. 2). The correlation was strongest (i.e. highest 265 266 Kendall's tau) for individual trees in the same stand (tau=0.51, Fig. 2c). We found yet stronger 267 correlations when we narrowed down the analysis on tree records occurring in Europe which we here define as those with a longitude smaller than 69°W (Fig. S2). Correlations would have been 268 269 considerably weaker if we had not corrected for age of the trees (Fig. S3). No phylogenetic 270 conservatism was found in either the variance of biomass fractions (Blomberg's K in the 271 cumulative analysis of the *Biomass-tree* data set was 0.37 with P = 0.95) or the distribution of 272 the species (Blomberg's K in the same data set was 0.53 with P = 0.12; Appendix 3). The 273 correlations were weaker after correction for phylogeny with phylogenetically independent 274 contrasts, except for the Biomass-tree data set (Fig. 3). The variances of the BAFs did not differ 275 significantly between angiosperms and gymnosperms in any of the three data sets (Fig. 4 – 276 evidence against hypothesis three).

The analysis using root:shoot ratios produced comparable results but with lower statistical power (Fig. 5a,b). The results of our analysis of the biomass data from America agreed with those for European trees (Fig. 5c; tau=0.91, *P*=0.07). Most of the trees in the datasets were classified as native and there were minimal differences in the additional analyses targeting natives (Fig. S4). There was insufficient statistical power to reach robust conclusions for invasives (Fig. S4).

283

284 *The influence of latitude*

285 There was no relationship between mean latitude and the variance of BAFs between foliage over 286 woody aboveground biomass in the Biomass-tree data set (Fig. 6a). The five tree species that 287 were observed both North and South of the latitudinal threshold of 58.5°N did not display any 288 patterns with regards to their BAFs (Fig. 6a – insert). *Picea obovata* showed higher BAFs in the 289 southern extent of its distribution, Pinus sylvestris and Betula alba in the northern extent of their 290 distribution whereas Larix sukaczewii and Pinus sibirica showed relatively balanced BAFs in 291 both extents. The lack of a relationship was even more apparent when we narrowed down our 292 analysis to individual stands in the *Biomass-tree* data set (data not shown). By contrast we could 293 observe a strong relationship between mean latitude and the variance of BAFs in the Biomass-294 plot data set (Fig. 6b).

295

296 Relative effect sizes of genetic and environmental variability

We had 40 observations and a mean slope of 3.324.10⁻⁵ per thousand km² of distribution in the 297 298 case of the individual stands in the *Biomass-tree* data set assessing genetic variability and 18 observations and a mean slope of 3.05. 10^{-2} per degree latitudinal range in the case of the 299 300 Biomass-plot data set measuring environmental variability. The resulting expression of variance (Appendix 4) was $s_p = 10^{-4} \sqrt{(0.11D^2 + 93,025\phi^2)}$ with D representing distribution range and φ 301 degrees latitudinal range. For average European settings of $D \approx 10,000$ thousands km² (also 302 consistent with (Tkach *et al.*, 2008)) and φ differences of 58.9° – 39.5° latitudinal range the two 303 factors inside the parenthesis take the values 11.10^6 and 35.10^6 suggesting that latitudinal range 304 305 (i.e. here used as a proxy of environmental variability) exerts on average a 48% stronger (i.e. the 306 resulting variances for genetic and environmental variability, if that the other factor was zero, 307 would be 0.33 and 0.59, respectively, giving relative proportions of 36% and 64%) influence on 308 variance in logRR of BAFs than distribution (i.e. here genetic variability) does. Genetic diversity 309 and latitudinal range, both induce strong changes on the variance in BAFs and the respective 310 standard deviations were multifold higher for species with extensive distributions (Fig. 2c) and 311 high latitudes (Fig. 6b).

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314 Discussion

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316 Living organisms are constantly challenged to optimally allocate their finite resources to 317 maximise fitness. This challenge leads to multiple investment trade-offs, many of which have 318 been extensively studied (Poorter et al., 2006; Cadotte, 2007; Huot et al., 2014). BAFs represent 319 direct expressions of some of these trade-offs and the phenotypes of the organisms. We 320 experimented with two such BAFs, foliage over woody aboveground biomass and the root:shoot 321 ratio, to show that both these fractions were more variable in tree species with extensive 322 distributions, even after limiting our analysis to neighbouring trees (Fig. 2c), in agreement with 323 hypothesis two. Our results were also valid after correcting for phylogenetic relatedness 324 (Appendix 3). A higher ability to adapt to a changing environment is one of the implications of 325 higher trait variance (Bolnick et al., 2011). The causality of this relationship is unclear and could 326 imply either that tree species that experience a higher phenotypic variability tend to have larger 327 distributions (e.g. van Vallen, 1973) or that some species are phenotypically more diverse in 328 response to a larger distribution and most likely a larger effective population size. Environmental 329 heterogeneity has been identified as a mechanism that facilitates genetic variation in plants 330 (Delph & Kelly, 2013), and a larger distribution implies a higher environmental heterogeneity. 331 Many tree species have extensive distributions, but individual trees can also disperse over very 332 large distances (Bacles et al., 2006; Petit & Hampe, 2006; Kremer et al., 2012) and cross-fertilise 333 with individuals experiencing differing environmental conditions. We thus favour the 334 interpretation that a larger distribution most likely induces a higher variance in BAFs. 335 BAFs in plants represent expressions of an equilibrium between stabilising 336 intrinsic/genetic factors such as ontogeny and the destabilising influence of the environment 337 (Barthélémy & Caraglio, 2007), i.e. any biomass-allocation fraction is an aggregate of these two

338 mechanisms. We thus present an argument detailing why our results were likely due to a higher 339 genetic variability of trees with more extensive distributions and not because of the 340 environmental conditions that the trees experienced (in support of hypothesis two and against 341 hypothesis one). The influence of the environment is expected to be only moderately important 342 at sites close to the centres of species distributions (compared to the edges). Most of the trees in 343 our data sets had been harvested near the centres of their distributions, because moderate 344 environmental conditions facilitate silvicultural practices. Most importantly, our results were 345 valid when using individual sites per species in which case the differences were most likely due to rapid evolution (i.e. the absence of systematic differences between angiosperms and

347 gymnosperms provided evidence against hypothesis three). Co-occurring tree species experience
348 comparable environmental conditions, so the resulting phenotypic variance should best represent
349 the genetic diversity (Fig. 2c).

350 Does latitude influence BAFs? We found no relationship in the Biomass-tree data set but 351 there was a strong positive relationship in the *Biomass-plot* data set (Fig. 6). A reason why the 352 results across the two data sets were incongruent probably relates to the way these were 353 standardized. The *Biomass-tree* data set contained multiple tree records per sampling site and 354 there were few replicate sites per tree species. Co-occurring woody species are likely to 355 experience comparable environmental conditions and the resulting BAFs should mainly reflect 356 genetic diversity. Because of the few replicate sites per tree species any influences of latitude 357 were masked in the data set by those of genetic diversity. This was even more pronounced when 358 we narrowed the analysis to a single site per tree species. By contrast in the *Biomass-plot* data set 359 via being limited to a single record per sampling site, we could effectively investigate differences 360 arising from environmental variability. Based on this analysis, trees that typically occur closer to 361 the poles exhibit higher variances of BAFs. Even though it has been disputed in the past 362 (Vázquez & Stevens, 2004), it is widely appreciated that environmental variability increases with 363 latitudinal range (MacArthur, 1972). This result is in accordance with our first hypothesis that 364 environmental variability should increase BAFs variance (Fig. 1a). The effect size we estimated 365 for environmental variability exceeded that for genetic diversity.

366 The survival challenges of tree species to the accelerating pace of global change is a key 367 topic in the biology of global change (Lenoir et al., 2008; Bertrand et al., 2016; Veresoglou & 368 Halley, 2018). Identifying tree species at a high risk of extinction is important. Several traits such 369 as longevity (Morin & Thuiller, 2009) and seed size (Veresoglou & Halley, 2018) might be 370 informative in terms of tree susceptibility to extinction. The loss of habitat is a key factor that 371 drives the eventual risk of extinction, but the relationship between habitat size and genetic 372 variability is poorly understood. The loss of genetic diversity following habitat loss can further 373 limit the ability of a species to cope with environmental conditions (Sexton et al., 2009) and 374 eventually accelerate extinction. The implications of extensive distributions have been debated. 375 A review of 31 studies by Lowe et al. (Lowe et al., 2005) found that habitat loss usually did not 376 significantly affect the genetic variability of tree species. Another meta-analysis by Vranckx et

al. (2012), however, found that habitat loss induced losses in the genetic diversity of species. An
extensive synthesis by Morueta-Holme et al. (2013) reported that plant species with extensive
geographic ranges were more genetically variable, and Kremer et al. (Kremer *et al.*, 2012)
argued that long-distance gene flow amongst trees likely conferred an evolutionary advantage. In
contrast, the implications of habitat loss in trees may be fully reversible if the former habitat of
these species can be restored before extinction (Newmark *et al.*, 2017).

383 Our results suggest that it is tree genetic diversity that induces variance in BAFs (Fig. 1b; 384 we found higher variance in BAFs in species with large distributions). Our analysis makes the 385 assumption that genetic variability is a good proxy of genetic diversity, which despite being a 386 common and well supported assumption in the literature (e.g. Avolio et al., 2012; Jöqvist and 387 Kremp, 2016), remains less robust than direct estimates of genetic diversity. Most importantly, 388 our general models did not discriminate between native and non-native plant species and used 389 real distributions to assess effect sizes. We often observe that the distributions of non-native (i.e. 390 invasive) species are not at equilibrium with their environment and that they possess a lower than 391 expected genetic diversity (Beaumont et al., 2009; Bradley et al., 2010). We would have 392 expected, as a result, relatively weak relationships between genetic diversity and realized 393 distributions for the subset of non-native trees which was not the case (Fig. S4). This finding 394 implies that changes in genetic diversity can occur quickly and develop after a few generations 395 of growth in isolation. Genetic diversity should thus play a key role in regulating the response of 396 species to future climate change, because of the extensive fitness implications of BAFs on the 397 ability of a species to adapt. The loss of habitat, even if transient, could lead to a loss of genetic 398 diversity, which would hinder species survival.

399 Human activity has sped up gene flow in almost all types of ecosystems and this should 400 continue in the near future (Wilson et al. 2009). On the short term, assisted gene flow 401 homogenizes populations (i.e. and their genetic diversity), allowing species to more effectively 402 cope with unfavourable environmental conditions (e.g. through acquiring more favourable 403 BAFs). Assisted gene flow should thus steepen the positive relationship between variance in 404 BAFs and distribution (Fig. 1b), benefiting disproportionally species with a large distribution at 405 the risk of species that maintain a limited genetic diversity. It has been documented, for example, 406 that invasions, which represent an alternative form of introductions of species with a high 407 competitive ability, induce extinctions of native species (Colautti et al. 2017; Catford et al.

408 2018). Even though we did not directly model assisted gene flow (as for example in Adams et al. 409 1998), the strong relationship between distribution and genetic diversity should be instructive for 410 forest management. Silvicultural practices such as the artificial regeneration of stands from 411 commercial genetic material (e.g. Bradshaw 2004; Finkeldey Ziehe 2004) might, therefore, be 412 precarious for the native species diversity. 413 In summary, BAFs were more variable in trees with extensive distributions, i.e. our 414 results supported only the second of our hypotheses. Most notably, we made the point that it was 415 a higher genetic variability that resulted in more variable BAFs for tree species with extensive 416 distributions. We thus present evidence that the loss of habitat for tree species through a quick 417 loss of genetic diversity could lower the ability of a species to modify its architecture (BAFs) in 418 response to changes in the environment (Fig. 1). It is thus likely that any loss of habitat may not 419 be as reversible as many believe. 420 421 Acknowledgements 422 We want to thank two anonymous reviewers and the editor for constructive criticism on our 423 paper. JP was funded by the European Research Council Synergy grant ERC-SyG-2013-610028 424 IMBALANCE-P. 425 426 Contributions 427 Conceived the study and carried out the analysis: SDV. The two authors wrote together the 428 article and approved the final version. 429 430 431 Data availability All the data we analyse here are already in the public domain. We present intermediate data in 432 433 our analysis in the form of Appendix 8. 434 435 436 Competing interests 437 The authors declare no conflicts of interest. 438

- 439 Supplementary Information
- 440 Fig. S1: Sensitivity analysis on the inclusion criterion of five full records
- 441 Fig. S2:Relationships between the variance in BAFs in records from Europe
- 442 Fig. S3: Weighted regressions between the variance in BAFs and the distribution of the species
- 443 without correcting for age
- 444 **Fig. S4**: Senstivity analysis on how relatioships differed between native and invasive tree species
- 445 **Table S1:** List of the plant species that we used in our analyses.
- 446 **Table S2**: Overview of all analyses.
- 447 **Table S3**: : Classification of tree species into native vs invasive
- 448 Appendix 1: Inclusion Criteria
- 449 Appendix 2: Effect sizes and model specifications
- 450 **Appendix 3**: Strength of BAF relationships
- 451 **Appendix 4**: Partitioning of BAFs into environmental and genetic variability
- 452 **Appendix 5**: Biases due to differences in coverage of the databases
- 453 **Appendix 6**: Phylogenetic tree
- 454 Appendix 7: Supplementary references
- 455 Appendix 8: Raw data
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642

643 Legends to Figures

644

645 Fig. 1 Conceptual diagram illustrating the three hypotheses we address in the manuscript. Four 646 hypothetical species, the broadleaves blue triangle and yellow circle (represented with an 647 elliptical leaf) and the conifers red rhomb and green square (represented by an acicular 648 leaf) are each sampled at four locations (map on top; note that distribution envelopes 649 differ). The ranking of their distributions is as follows: red rhomb, blue triangle, yellow 650 circle and green square. In Hypothesis One (a) we expect that the fraction of variance in 651 biomass allocation fractions that is explained by genetic variability (purple discontinuous 652 lines) is independent of distribution and that environmental factors increase the variance 653 (green arrow) more in those species that have a small distribution (larger arrow) resulting 654 in a negative relationship between observed variance and distribution (green 655 discontinuous lines). In Hypothesis Two (b) we expect that a larger distribution results in 656 a higher genetic diversity which is depicted with a purple line. The environment increases 657 variance irrespective of distribution (green arrows) resulting in a positive relationship 658 between observed variance and distribution (green discontinuous lines). In Hypothesis 659 Three (c) the biomass allocation fractions vary independently of distribution and can be 660 explained by the evolutionary history of the plant (here angiosperms vs gymnosperms). We could assess genetic variance in BAFs by comparing conspecific trees in the same 661 662 stand after correcting for age-differences and the sum of genetic and environmental 663 variability by comparing across stands, also after correcting for age. 664

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666 Fig. 2 Relationships between the variance in biomass-allocation fractions (foliage over woody 667 aboveground biomass) and distribution of the tree species. (a) Biomass-tree data set, all 668 possible observations (multiple trees per location) of the specific data set (lower than in 669 Biomass-plot); (b) *Biomass-plot* data set, one observation per location (at a plot level); (c) 670 Biomass-tree data set, observations per species only describe trees in the stand (plot) that 671 contained the most tree observations. The dashed lines represent the best fits. Numbers 672 next to the data points indicate the number of observations per species used to calculate 673 the variance. Overlaid map shows in red the location of the sites where the data

| 674 | originated from - we only analysed woody species with an extensive distribution (over 2 |
|-----|--|
| 675 | million km ²) in Europe. Phylogenetic correction was not applied. Relationships were |
| 676 | stronger when we repeated the analyses with the subset of sites located in Europe (Fig. |
| 677 | S2). |
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| 682 | Fig. 3 Relationships between variance in biomass-allocation fractions (foliage over woody |
| 683 | aboveground biomass) and distribution of the tree species for phylogenetically corrected |
| 684 | data using phylogenetically independent contrasts (PIC). (a) Biomass-tree data set, all |
| 685 | possible observations; (b) Biomass-plot data set, one observation per location; (c) |
| 686 | Biomass-tree data set, observations per species are from the stand that contained most |
| 687 | observations. The dashed red lines represent the best fits. The overlaid phylogenetic tree |
| 688 | depicts in the form of squares information (blue: first quartile; white middle two |
| 689 | quartiles; red 4th quartile) on variance in biomass allocation fractions (left) and |
| 690 | distribution (right) of the woody species included in the analysis and their phylogenetic |
| 691 | relationships (tree). Note the absence of phylogenetic signal which was confirmed with |
| 692 | Blomberg K tests. |
| 693 | |
| 694 | Fig. 4 Differences in variance in biomass-allocation fractions (foliage over woody aboveground |
| 695 | biomass) between angiosperms (elliptical leaf in yellow) and gymnosperms (acicular leaf |
| 696 | in green). (a) Biomass-tree data set, all possible observations; (b) Biomass-plot data set, |
| 697 | one observation per location; (c) Biomass-tree data set, observations per species are from |
| 698 | the stand that contained most observations. None of the t-tests were significant. |
| 699 | |
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| 701 | Fig. 5: (a; b) Relationships between the variance in biomass-allocation fractions (root over shoot |
| 702 | fractions) and distribution of the tree species. (a) Biomass-tree data set, all possible |
| 703 | observations; (b) Biomass-tree data set, observations per species are from the stand that |
| 704 | contained most observations; (c) variance in biomass-allocation ratios for Pinus echinata |
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(observed in 23 states), *Pinus taeda* (observed in 20 states, from two studies) and *Pinus palustris* (observed in 10 states) across four studies in the Legacy Tree Data database
(USA). The three species have nested distributions in America. The dashed lines
represent the best fits. Numbers next to the data points indicate the number of
observations per species used to calculate the variance. Overlaid map shows in red the
location of the sites where the data originated from.

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713 Fig. 6 Influence of latitude on variance in biomass-allocation fractions (foliage over woody 714 biomass). The x-axis describes the average latitude of the distribution range of the woody 715 species whereas the y-axis the variance in biomass-allocation fractions. Panel (a) 716 describes the Biomass-tree data set whereas panel (b) the Biomass-plot data set. The 717 insert in panel (a) shows how variance differs in the north and south distribution of five 718 tree species in the Biomass-tree data set (opaque triangles). North here describes latitudes 719 in Europe higher than 58.5°N (mean of the latitudinal extremes of European territory) and 720 South lower than this value. These five species were the only ones that met the inclusion 721 criterion, namely being represented with a minimum of five entries both in North and 722 South Europe. The continuous opaque line is an isocline where variance in the North 723 equals that in the South.

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Distribution

Distribution



Variance in biomass-allocation franctions







