

Production and preservation of the smallest drumlins

Journal:	GFF
Manuscript ID	SGFF-2017-0047.R1
Manuscript Type:	Special Issue
Date Submitted by the Author:	27-Feb-2018
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Keywords:	Drumlin, Iceland, Múlajökull, size-frequency distribution, landform mapping, subglacial bedform, morphometry

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17	11	
19	12	Acknowledgements
20 21	13	
22 23	14	We thank the organisers of the 'Beauty of Drumlins' symposium for bringing the authors
24 25	15	together in Lund, Sweden, May 2017, in honour of Professor Per Möller upon his retirement.
26	16	We are grateful to Chris Stokes and an anonymous reviewer for their insightful comments,
27 28	17	which improved the manuscript.
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22 Abstract

Few very small drumlins are typically mapped in previously glaciated landscapes, which might be an important signature of subglacial processes or an observational artefact. 143 newly emergent drumlins, recently sculpted by the Múlajökull glacier, have been mapped using high-resolution LiDAR and aerial photographs in addition to field surveying. In this paper, these are used as evidence that few small drumlins (e.g. height $H \leq 4$ m, width $W \leq 40$ m, length $L \leq 100$ m) are produced; at least, few survive to pass outside the ice margin in this actively forming drumlin field. Specifically, the lack of a multitude of small features seen in other landforms (e.g. volcanoes) is argued not to be due to i) Digital Elevation Model (DEM) resolution or quality, ii) mapper ability in complex (i.e. anthropogenically cluttered or vegetated) landscapes, or iii) post-glacial degradation at this site. So, whilst detection ability must still be at least acknowledged in drumlin mapping, and ideally corrected for in quantitative analyses, this observation can now be firmly taken as a constraint upon drumlin formation models (i.e. statistical, conceptual, or numerical ice flow). Our preferred explanation for the scarcity of small drumlins, at least at sites similar to Múlajökull (i.e. ice lobes with near-margin drumlin genesis), is that they form stochastically during multiple surge cycles, evolving from wide and gentle pre-existing undulations by increasing rapidly in amplitude before significant streamlining occurs.

44 Keywords

45 Drumlin, size_frequency_distribution, Múlajökull, Iceland, landform mapping, subglacial
46 bedform, morphometry.

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1. Introduction

Drumlins are subglacial bedforms aligned parallel to ice flow, created by interactions in the ice-sediment-water system underneath glaciers or ice-sheets [e.g. Menzies, 1979; Clark et al., 2009; Benn and Evans, 2010]. Their mode of formation remains enigmatic and debated [Smalley and Unwin, 1968; Menzies, 1979; Shaw, 1983; Boulton and Hindmarsh, 1987; Hindmarsh, 1998; Fowler, 2000], primarily because the bases of modern ice sheets are inaccessible, which results in few direct observations [King et al., 2007; Smith and Murray, 2009]. Mapped morphometrics of the numerous (i.e. \gg 10,000) drumlins formed during past glaciations [e.g. Hättestrand et al., 2004; Storrar and Stokes, 2007; MacLachlan and Eyles, 2013] are therefore key to understanding the subglacial interface, despite less readily yielding secure conclusions about the dynamics and mechanics of former ice sheets.

Observations of bedform position and morphology are used to indicate ice extent or flow direction [e.g. Hollingsworth, 1931; Livingstone et al., 2008], for example to assess consistency with numerical ice sheet models [Evans et al., 2009]. Elongated bedforms have also been linked to fast ice flow [Clark, 1993; Stokes and Clark, 2002]. However, it is rare to directly or quantitatively use bedform morphometrics to consider the mechanics of ice-sediment interaction and flow [e.g. Chorley, 1959; Smalley and Warburton, 1994]. As a step to bridging this gap Hillier et al. [2013] proposed a conceptual model to explain the size-distributions of subglacial bedforms in terms of stochastic ice-sediment-water interaction; subsequently, a variety of statistical models have been developed to formalize the postulated stochastic behaviour [Fowler et al., 2013; Hillier et al., 2016]. Hence, as a theoretical basis emerges for interrogating bedform size observations in more depth, high-quality morphometric data are becoming more important.

Drumlins have heights (H) (a.k.a. amplitude) ranging up to a few 10s of m, their widths (W) are of the order of 100s of m, and they have lengths (L) of up to a few km [e.g. Hollingsworth, 1931; Hättestrand et al., 2004; Clark et al., 2009]. Size distributions can be summarized by basic statistics (e.g. mean, minimum, maximum, modal class, skew) [e.g. Clark et al., 2009], or by one- or two-parameter functions (i.e. exponential, log-Normal, Gamma) [Fowler et al., 2013; Hillier et al., 2013, 2016] approximating a ubiquitous typical shape (Figure 1). There are few small bedforms mapped, a modal peak at sizes above this forming a 'roll-over', and an approximately exponential tail of frequencies decreasing towards the largest sizes. This is true

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81 for both aggregated data and, importantly, individual flow sets that likely represent 82 glaciological conditions at a particular location and time [*Hillier et al.*, 2013, 2016].

FIG 1 HERE

The roll-over and absence of very small forms might be an important signature of subglacial processes, or be due to post-glacial degradation, or just be an observational artefact (e.g. due to low DEM resolution) [see Hillier et al., 2013]. If this absence is real it could be a key constraint on drumlin formation, for instance distinguishing between statistical models built to represent various glaciologically reasonable hypotheses [Hillier et al., 2016]. Illustratively, for landslides a roll-over has been interpreted in terms of physical processes (e.g. cohesion contributing to soil stability) [e.g. Malamud et al., 2004; Frattini and Crosta, 2013] and, contrastingly, elsewhere considered as observational under-sampling [e.g. Stark and Hovius, 2001; Ten Brink, 2006]. Submarine volcanoes tend to have no roll-over demonstrating that natural processes can also produce sizes that can be approximated by simpler distributions such as exponential or power-law [e.g. Smith and Jordan, 1987; Scheirer and Macdonald, 1995; Rappaport et al., 1997; Hillier and Watts, 2007; Bohnensteihl et al., 2008].

In terms of drumlin mapping, Spagnolo et al. [2012] assert that a small drumlin of 2.1 m relief (*H* = 2.1, *W* = 150, *L* = 430) is reliably mapped in the 5 m resolution NEXTmap BritainTM InSAR-derived DEM. Indeed, a field visit is used to verify its existence. Using the same data product, Hillier et al. [2014] use synthetic landscapes to illustrate that as much as 75% of the smallest drumlins might be missed during mapping in complex landscapes (i.e. with anthropogenic clutter and trees), and that amplitude (or height) is the key variable governing detectability (Figure 2). The designed landscapes of Hillier et al. [2014] were Digital Elevation Models (DEMs) of a real glaciated area that had 173 drumlins of realistic morphology and size placed within them. Without prior sight of the drumlins' locations, 27 operators then mapped the area to assess their effectiveness. Their study site near Loch Lomond is challenging to map, so this perhaps illustrates a conservative 'worst case' for under-detection. Thus whilst a 'small' feature, defined here to be less than about half the modal size of a dataset, can be mapped, an open question remains as to the completeness of mapping at these small sizes and its impact on size-distributions and inferences from them.

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116 Individual drumlins, or small groups of drumlins, have been described under ice streams [King 117 et al., 2007; e.g. Smith et al., 2007] and in front of contemporary glaciers in Alaska [Haselton, 118 1966], Antarctica [e.g. Rabassa, 1987], Switzerland [van der Meer, 1983] and Iceland [e.g. 119 Boulton, 1987; Krüger, 1987; Evans and Twigg, 2002]. The number (i.e. 143) of drumlins in the 120 flow set at the Múlajökull surge-type piedmont glacier in central Iceland is large for a 121 contemporary glacier and, as yet, unique for a large and active drumlin field in being both 122 currently sub-aerial and the subject of detailed geomorphological, sedimentological and 123 stratigraphic analysis (Figure 3) [e.g. Johnson et al., 2010; Benediktsson et al., 2016]. It is therefore a study site with the power to yield novel insights, but differences across the site in 124 simple descriptive measures of the drumlin morphometrics (e.g. mean W) [e.g. Benediktsson et 125 126 al., 2016] have not yet been verified by statistical testing, nor have the size-frequency 127 distributions been investigated in detail.

129 In this paper, 143 newly emergent drumlins recently created by the Múlajökull glacier (Figure 3c) are used to understand the production and preservation of the smallest drumlins. They 130 have had little time to degrade post-glacially, have no 'clutter' on their surfaces (e.g. trees, 131 132 houses), and are mapped in high-resolution data supported by extensive ground-truthing during fieldwork [Benediktsson et al., 2016], removing many sources of observational ambiguity. 133 134 Error bars, statistical significances, and distribution parameters (i.e. exponential, log-Normal, 135 Gamma) are computed to robustly examine this Icelandic data, and drumlin mapping from the 136 UK and Sweden are used to put it in a wider context. Physically-based statistical models of drumlin formation [e.g. Hillier et al., 2016] are used to assist in interpreting the size-frequency 137 138 observations, the output of which is blended with field observations (e.g. sedimentology, stratigraphy) to offer a model that explains the scarcity of small drumlins at sites like Múlajökull 139 140 (i.e. ice lobes with near-margin drumlin genesis).

FIG 3 HERE

144 2. Study area

Múlajökull is a surge-type glacier in the southern part of the Hofsjökull ice cap in central Iceland
[*Björnsson and Pálsson*, 2008] with surges recorded in 1924, 1954, 1966, 1971, 1978-79, 1986,
1992, and 2008 [*Björnsson*, 2009]. The glacier forefield is relatively flat; it dips gently in a down-

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ice direction away from the glacier_at_approximately 1° [McCracken et al., 2016], is at roughly 600 m above sea level, and contains 132 fully-exposed and 11 partially-exposed drumlins of roughly elliptical planform interspersed with small lakes [Jónsson et al., 2014; Benediktsson et al., 2016]. Even more drumlins have been reported to be beneath Múlajökull's margin, but limited to a 0.5-0.7 km wide zone inside the 2015 ice margin [Lamsters et al., 2016]. Beyond this limit, farther from the ice margin, the GPR survey reveals no drumlins [e.g. Fig. 4 of Lamsters et al., 2016], and in the segment surveyed the bed starts dipping up ice into the prominent, ~130 m deep subglacial overdeepening located in the centre of the Múlajökull ice lobe [Björnsson, 1986]. The exposed drumlins comprise multiple till beds [e.g. Johnson et al., 2010]. As such, they are examples of mainly till-cored drumlins rather than other variants such as 'crag-and-tail' [Phillips et al., 2010; Stokes et al., 2011; Dowling et al., 2015]. Surface till terminating at the 1992 moraine, field evidence of stagnant ice that could not have deposited substantive thicknesses of till after the 1992 surge (e.g. preserved flutes), non-deposition of till during small winter advances, and till shear fabrics that conform to drumlin morphology all indicate that tills were deposited during surges [Johnson et al., 2010; McCracken et al., 2016]. The youngest till bed roughly replicates the drumlins' form and truncates stratigraphically lower units, particularly on the drumlins' flanks and heads [Johnson et al., 2010; Benediktsson et al., 2016]. This indicates that during surge-cycles drumlins likely get progressively narrower and higher [Benediktsson et al., 2016].

The drumlins' reported sizes [Johnson et al., 2010; Jónsson et al., 2014; Benediktsson et al., 2016] are similar to widespread and well-studied Pleistocene drumlin fields [Patterson and Hooke, 1995; Clark et al., 2009; Hillier et al., 2013]. Ice proximal drumlins are more elongate than distal ones [e.g. Benediktsson et al., 2016]. This study area, therefore, despite differences in spatial extent, perhaps most directly relates to Pleistocene drumlin fields where elongation ratio (i.e. L/W) increases up-ice, namely away from <u>a</u> margin related to <u>a</u> relevant maximum ice extent [e.g. Colgan and Mickelson, 1997; Stokes and Clark, 2003].

This drumlin field is argued to be 'active' in the sense that it is sculpted by the current glacial
regime of repeated surges and intervening quiescent phases [e.g. *McCracken et al.*, 2016], most
recently and directly evidenced by a till from the 2008 surge lying atop an erosional surface [i.e. *Johnson et al.*, 2010; *Benediktsson et al.*, 2016]. Subglacial morphological dynamics at any given
location may be punctuated by periods without change, and therefore be inactive at any exact

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instant, even under flowing ice-streams. Thus, 'active' does not refer to changes this minute or even today, but relates to present conditions and a time-scale is implied, in this case decades.

The foreland of Múlajökull is comprised of minimally vegetated and essentially homogenous till and outwash deposits [Jónsson et al., 2014; Benediktsson et al., 2016]. Specifically, there is no significant vegetation inside the Little Ice Age (LIA) moraine that bounds the immediate foreland and even though the vegetation cover on the moraine and beyond is continuous, it is limited to short grasses and shrubs under a few 10s of cm in height (Figure 4). There has been no anthropogenic disturbance (e.g. houses or infrastructure) in the area. Even drumlins proximal to the glacier are not ice-cored [Benediktsson et al., 2016]. There are no large topographic variations that might dominate ice flow patterns, such as the bedrock ridge near Lough Gara in Ireland [cf. Hillier and Smith, 2008]. So variations in drumlin morphology cannot be attributed to large-scale topography, preservation impacts of internal ice melting, or till type, and there is no evidence of bedrock variation.

The maximum Holocene extent of Múlajökull was reached in the LIA (1717-1758), recorded by the Arnarfellsmúlar terminal moraine [Benediktsson et al., 2015]. The most substantial surges since 1924 (i.e. 1954, 1971, 1986, 1992) have terminated approximately at the remaining 1992 end moraine [Björnsson et al., 2003; Johnson et al., 2010]. Also, a small surge in 2008 was observed to create a significant ice-cored moraine just distal of the present ice margin [Jónsson et al., 2014; Benediktsson et al., 2016]. As such, a series of moraines outside the 1992 limit, including an overridden moraine, inboard of the Arnarfellsmúlar terminal moraine suggest that this area also experienced multiple surges during the LIA both before and after the maximum extent in the early to mid-1700s [Jónsson et al., 2014; Benediktsson et al., 2015]. Thus, it is convenient to divide the forefield into two zones 'inside' and 'outside' the 1992 moraine based on historical surge activity. The area inside is reported to contain more elongate_drumlins than outside, with respective mean elongation ratios (i.e. L/W) of 3.0 and 1.9 [Benediktsson et al., 2016]. It has been hypothesized [Johnson et al., 2010; Jónsson et al., 2014; Benediktsson et al., 2016] that distal drumlins have been shaped by fewer surges than those closer to the glacier. At Múlajökull surges deposit till with a sedimentology and stratigraphy that imply net aggradation [Johnson et al., 2010; McCracken et al., 2016], so inferred thicker proglacial sediment near to the current ice margin implies more geomorphically active surges there [McCracken et al., 2016]. This inference is supported by a number of lines of evidence. Topography dips away from the glacier aligned with flow parallel features (e.g. flutes) and perpendicular to terminal

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moraines, indicating that it reflects ice flow rather than other controls. A break in slope exists at the 1992 moraine, where the four most recent large surges have stopped. There is no evidence for bedrock control of slope, and if it is postulated to be causing up-ice dips at this site its influence is demonstrated to be subservient to ice flow by the overdeepening just upstream of the current ice margin [Björnsson, 1986; Lamsters et al., 2016]. Thus, a powerful aspect of the Múlajökull site is that relatively strong constraints exist on the timing and duration of geomorphic work in two zones, which is rare. This constraint allows predictions by models of how subglacial bedforms (e.g. drumlins) progressively evolve with time to be considered against observations that have quite low levels of ambiguity.

Neither sedimentology nor stratigraphy yet directly constrain drumlins' elongation during surge-cycles. Till fabrics and bulk densities indicate that inter-drumlin areas have experienced higher maximum effective stresses (~100 kPa), argued to represent guiescent periods under the assumption of effective and channelized drainage at these times [McCracken et al., 2016]. Then, as in other models [e.g. Hindmarsh, 1998; Chapwanya et al., 2011], increased effective shear stresses are taken to indicate higher rates of sediment transport. The basal stress distribution is asserted to be compatible with a crevasse pattern at the ice front, which is strongly related to the spatial pattern of the drumlins [Johnson et al., 2010; Benediktsson et al., 2016; McCracken et al., 2016], but the mechanics of causal relationship remain conjectural. The available observations have been consolidated and reconciled into a conceptual model [Johnson et al., 2010; Jónsson et al., 2014; Benediktsson et al., 2016; McCracken et al., 2016], an extreme precis of which follows: Although sediment transport mechanisms are not uniquely constrained, surges deposit drapes of till everywhere, then in each intervening quiescent period there is erosion in the inter-drumlin areas, processes that combine to lead to increases in H and L but a decrease in W. A mathematical model has been developed to formalize this [Iverson et al., 2017]. GPR data [Lamsters et al., 2016] and the Múlajökull drumlins' proximity to the LIA terminal moraine dictate that these models are based on, and therefore most directly constrain, near-margin (i.e. < 1-2 km) drumlin formation.

FIG 4 HERE

247 3. Data and mapping

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Johnson et al. [2010] use 1 m resolution gridding of airborne laser scanning (i.e. LiDAR) data from 2008 to map >50 drumlins inside the 1992 moraine, which are 90-320 m long (\overline{L} = 185 m), 30-105 m wide (\overline{W} = 64 m), and 5-10 m in relief. Mean elongation ratio is 3.0. Drumlins outside the 1992 moraine were first mapped by Jónsson et al [2014] from a 3 m resolution DEM that was created from aerial stereophotographs taken in 1995, increasing the total number of drumlins to 110. The size ranges increased, e.g. W values are 20-180 m, as is expected of a larger sample. Most recently, 143 drumlins were mapped from 0.5 m resolution LiDAR collected in 2013 [Benediktsson et al., 2016] (see Supplementary Material), reporting broadly comparable morphometrics, which are similar to widespread and well-studied Pleistocene drumlin fields [Patterson and Hooke, 1995; Ó Cofaigh et al., 2010; Hillier et al., 2013]. The existence and conformity to expectations of shape of all drumlins mapped were verified by inspection in the field, and no additional small drumlins were identified whilst on the ground [Johnson et al., 2010; Jónsson et al., 2014; Benediktsson et al., 2016].

A first source of uncertainty in the drumlin morphometrics at Múlajökull might be DEM resolution or quality. Whilst the 1995 DEM is based on stereophotogrammetry and is of a low resolution and accuracy [Jónsson et al., 2014], LiDAR data are widely regarded as a good basis for producing high quality DEMs in glacial and pro-glacial areas [e.g. Favey et al., 1999]. Even the 2008 LiDAR data have a point density of 0.33 m^{-2} , an average of 10 data per 5x5 m grid cell, and estimated horizontal accuracy of <0.5 m [Jóhannesson et al., 2014] much below drumlins' planform dimensions (i.e. L, W). This assertion is supported by close agreement (i.e. mean vertical difference of 0.132 m) between the 2008 and 2013 LiDAR DEMs on two selected profiles (Figure 5). Of particular interest is the inter-survey agreement between the size of the undulations shown, with variance in amplitude on the order of 0.1 m, which is less than the H of even the smallest mapped drumlins. Thus, uncertainty in drumlin morphometrics from DEM creation will be small. Accurate DEM creation is, at least in part, due to minimal vegetation cover. Vegetation present within the area of the 2013 LiDAR DEM is mainly in the form of mosses that are limited to streams, shallow ponds and wet ground (Figure 4), and will typically be penetrated by the LiDAR sensing method.

A second potential discrepancy between mapped drumlins and the population of subglacially produced forms they preserve and reflect is post-glacial alteration. Plan view comparison

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(Figure 6) indicates that surface alteration in the foreland between 2008 and 2013 is relatively small (i.e. < 0.5 m, dark green) with respect to drumlin dimensions. Exceptions to this are readily explicable, namely the lowering of the level of ice-marginal lakes by 0.5-2 m [Benediktsson et al., 2016] and the degradation of the 2008 ice-cored moraine, which is superimposed on some drumlins. In terms of ground-truthing, terrain profiles surveyed using a TopCon GTS-236N total station between 2011 and 2014 show 0.4-0.8 m lowering of the 2008 ice-cored moraine crest due to melting, but negligible (i.e. <0.1 m) surface alteration on drumlin surfaces outside the degrading moraine. A similar conclusion is reached by comparing change between 2008 and 2013 for 'high' and 'low' stretches of profiles P1 and P2 extracted from the LiDAR data in locations shown on Figure 5. In P1 high and low areas changed by +0.073 and +0.055 m, respectively, apparently indicating an amplitude increase of ~ 2 cm, which is inconsistent with gravity driven mass-wasting. In P2, an amplitude decrease of ~1 cm is implied. Taken together, minimal change at the limit of observational resolution is demonstrated. This is consistent with stability at bedform scales (i.e. few 100s of m) over decades observed in other Icelandic till plains such as at Brúarjökull [Korsgaard et al., 2015]. Thus, there is no evidence of post-glacial degradation substantively impacting drumlins at this site.

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FIG 6 HERE

A third source of disagreement has the potential to arise in the methods used in mapping [Podwysocki et al., 1975; Siegal and Short, 1977; Smith and Clark, 2005; Gardin et al., 2011; Ardelean et al., 2013; Van Coillie et al., 2014] and then calculating the metrics of drumlins [e.g. Spagnolo et al., 2012; Hillier and Smith, 2014; Jorge and Brennand, 2017]. During fieldwork [e.g. Jónsson et al., 2014] an ambiguity of 1-10 m was identified in setting the location of boundaries that were not at shorelines, making a GIS approach [e.g. see Smith et al., 2006; Spagnolo et al., 2012] the most consistent and reproducible way of delimiting drumlins. Conceptual ambiguity exists where lakes conceal the land surface, with a debate as to whether drumlins are best defined as isolated features or waveforms [Stokes et al., 2013b], but no DEM mapping approach is a solution for this. Benediktsson et al. [2016] identified and mapped the drumlins at a scale of 1:3000-1:6000 in ArcGIS 10.2.2 using a hillshade model of the 0.5 m LiDAR DEM from 2013 with 1.5-4 times vertical exaggeration, 20-30° solar angle and illumination azimuths at 45° and 315°. A combination of slope analysis and visual inspection was used to delimit the drumlins at a break in slope, which could either be abrupt (e.g. lakes, outwash) or gradual. Only where both axes of a putative drumlin (long, short) were upstanding from the landscape was

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the planform shape considered delineated and a drumlin defined [Dowling, 2016]. However, if small forms have similar morphology to their larger companions, this will not impact the size-frequency distribution of drumlins (see Section 6.1).

Drumlin length (L) and width (W) were derived by measuring the length of the longest lines parallel and perpendicular to ice flow, respectively, within each drumlin [Benediktsson et al., 2016]. Being recently deglaciated, ice-flow direction was determined by flutes and other streamlining in the forefield. Drumlin relief (H) was defined by the range in elevation of each drumlin. Whilst this is a simple approach [e.g. see Hillier and Smith, 2012, 2014; Spagnolo et al., 2012], lakes bounding all but 3 drumlins inside the 1992 moraine make a horizontal basal plane a natural choice, with consistency requiring the same to be done outside the 1992 moraine. In this particular site, the use of drumlin elevation range to represent drumlin height (i.e. H) will cause relatively minor artefacts since the slope of the foreland is shallow (i.e. 0.007-0.023) [McCracken et al., 2016]. At least, the shape of the size distributions for H (Figure 7) will likely be minimally affected (see Section 4.4).

A fourth and final source of ambiguity in drumlin morphometrics inside the 1992 moraine is the presence of pro-glacial lakes. Draining these may increase mapped estimates of H, L, and W for lake-bounded features. The magnitude of this is difficult to constrain without additional information as it will depend on drumlin shape in the flooded areas, which is currently unknown, and is possibly influenced by lake drainage itself. However, the scarcity of many small drumlins between larger ones mapped outside the 1992 moraine indicates that this ambiguity will not impact the existence or otherwise of a roll-over in size frequency distributions.

4. Statistical analysis of the Múlajökull drumlin dataset

The median, range (i.e. minimum and maximum), standard deviation, and mean of drumlin sizes have been reported for Múlajökull [Benediktsson et al., 2016], but the uncertainty associated with size measurements has not. Here selected error bars, statistical significances, and distribution parameters (i.e. exponential, log-Normal, Gamma) are computed. Twoparameter distributions are a more sophisticated description of size-distributions [e.g. Hillier et al., 2013; Ely et al., 2017], and are reported for use in future compilations and analysis (Table 1). The statistical analysis also allows a robust evaluation of sizes observed at Múlajökull (Section 6). In particular, Section 4.1 reports error bars and sensitivity tests for the measures of central

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tendency (i.e. mean of H, L, and W) that are a key part of the design of the conceptual model at Múlajökull [i.e. Johnson et al., 2010; McCracken et al., 2016; Iverson et al., 2017]. Sensitivity to the measure of central tendency used (i.e. mean, median or mode) is also considered. Section 4.2 determines if a roll-over exists, thereby permitting comment on whether or not small drumlins are scarce. To test posited models of drumlin formation that have been statistically formalized [Fowler et al., 2013; i.e. Hillier et al., 2013] Section 4.3 summarizes key variations in the relevant metrics of the two-parameter distribution. Finally, Section 4.4 considers a potential systematic (i.e. not due to randomness and sampling) issue with the method used to calculate

359 <u>H.</u>

<u>4.1 Robustness of Observations</u>

Given the relatively small number of data inside and outside of the 1992 moraine (i.e. 77 and 55) it is <u>possible that an</u> apparent trend or observation <u>does not</u> actually exist (i.e. is <u>not</u> statistically significant). It could <u>arise simply</u> due to random variation in the selection of a sample; conceptually, observed data are a sample reflecting a parent population of what could be produced under identical glaciological conditions. A convention of considering a 5% chance of the result occurring by random variation in sampling (i.e. 95% level, p < 0.05) is arbitrary [*Wasserstein and Lazar*, 2016], so both 5% and 10% levels are reported.

The mean (i.e. μ) is perhaps the most commonly computed statistic for this type of <u>bedform</u> data, and relates to the Normal distribution (Table 1). A Welch t-test (2-tailed) confirms the observations of Benediktsson *et al.* [2016] that drumlins inside the 1992 moraine are longer (*P* \ll 0.01) and narrower (*P* = 0.011) than those outside, and strengthens their view that there appears to be no increase in heights.

All of the dimensions (i.e. *H*, *W*, *L*) for both data sets appear positively skewed, indicating distributions with <u>a more heavily populated right-hand</u> tail t<u>han</u> a Normal distribution, and most <u>of the skews (4 of 6)</u> are statistically significant (P < 0.05). <u>In other words, the distributions</u> <u>are not Normal, so a different measure of central tendency may be a more appropriate</u> indicator of where the distribution is located on the *x*-axis if plotted (e.g. Figure 1a). However, there is consistency (Table 1) between measures of central tendency (i.e. mean, median and <u>mode) alleviating any concern about previous uses of the mean</u>.

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4.2 Existence of a roll-over

To determine the existence, or otherwise, of a roll-over the exponential [Hillier et al., 2013], log-Normal [Fowler et al., 2013; Hillier et al., 2016] and Gamma [Hillier et al., 2016] distributions are fitted to the observations (Figure 7). Reassuringly, the variations in parameters of these distributions (μ_l , σ_l , α , β and λ) also show broad consistency with change in μ between the zones inside and outside the 1992 moraine. Namely, between the two areas H is similar and L is different, with a weaker signal for W reflecting a smaller magnitude of change. Both log-Normal and Gamma distributions fit the data comparably well, whilst the exponential fits the upper tail (i.e. larger forms) only and not the roll-over. In other words, there are fewer small drumlins than expected by simple extrapolation (i.e. using the exponential) from larger drumlins, which are typically more reliably observed than smaller drumlins. For completeness, a conservative correction for under-sampling to bend the exponential model towards the data (i.e. potentially account for the roll-over) is shown (dashed grey line) to allow for a direct comparison with analyses for other areas (see Section 5). Even were this correction for cluttered, hilly, 5 m resolution, InSAR-derived data applicable, it is insufficient to explain the roll-over; a factor of x10 (i.e. 10% recovery) equates to 2.3 on the vertical scale, which is a natural logarithm. In short, a roll-over exists in the data from Múlajökull, indicating a scarcity of small drumlins either side of the 1992 moraine.

4.3 Quantities Relating to Statistical Models

Distribution parameters that can be related to physically-based statistical models of drumlin formation are μ_L , σ_L , α , β and λ [Hillier et al., 2016]. The models predict how these parameters will change as the Múlajökull area evolved geomorphologically, a time-progression represented by the difference between drumlins outside as compared to inside the 1992 moraine. Although individual changes should be treated with caution where they are not statistically significant, patterns or trends across multiple dimensions (i.e. H, W, L) might not be coincidental; illustratively, if an observation about H and W both agree with the model but each with a 20% chance of occurring through random variation (i.e. P = 0.2), then the chance of them both occurring randomly is only 4% (i.e. P = 0.2*0.2 = 0.04). In Table 1 α consistently decreases as <u>drumlins evolve</u>, as does β in H and L. Unsurprisingly, being a similar quantity (i.e. rate in the statistical models), λ follows the same pattern as β , except there some statistical significance for the increase in W even when considered in isolation. As required, being a very similar

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419 quantity, μ_{L} behaves as μ_{i} *L* increases as drumlins evolve, *W* decreases and *H* is apparently 420 roughly stable. Whilst not readily understandable out of context, these variations invalidate one 421 of the two main physically-based statistical models of drumlin formation [Hillier et al., 2016] 422 (see Section 6.2.3).

4.4 Impact of Relief Quantification Method

Benediktsson et al. [2016] use range to quantify drumlin height, a method that has been called into question [e.g. Smith et al., 2009; Hillier and Smith, 2012; Spagnolo et al., 2012]. To approximately assess the impact of a more sophisticated H computation [e.g. Hillier and Smith, 2012], a correction of $H_c = H - (L/2)^*g$ is applied, where g is slope. 0.01 is an approximate central value for the slope of the foreland, which varies between roughly 0.007 and 0.023 [see McCracken et al., 2016]._Slopes of the fitted trends from P1 and P2 are 0.89° and 0.38°, respectively. The dip direction aligns with drumlin elongation. The lowest point, used in the vertical range quantification, will typically be near the drumlin's distal end, whilst the highest will be somewhat central. Thus, removing the slope between the centre and edge of each drumlin approximates the overestimation of H when range is used [e.g. see Spagnolo et al., 2012]. After applying the correction, the shape of curves equivalent to those in Figure 7a,d is not substantively or visibly different, supporting the idea that using range [Benediktsson et al., 2016] is insufficient to invalidate the conclusions reached here about the scarcity of small drumlins. Indeed, even correcting for more detailed effects, such as applying a different slope either side of the 1992 moraine [McCracken et al., 2016] cannot alter the shape of a sizefrequency distribution further. With the correction \overline{H} becomes 6.8±0.3 (2 s.f.) both inside and outside the 1992 moraine, with no significant difference (p = 0.942). Using a different slope either side of the 1992 moraine, however, may alter mean values (e.g. Section 4.1), even if such complexity in a correction likely exceeds the validity of the approximate assessment used here. To be safe, a full re-computation of *H* values is recommended in future.

451 5. Statistical analysis of UK and Swedish drumlin datasets

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FIG 7 HERE

TABLE 1 HERE

To place the observations at Múlajökull in a wider context, the extent to which they are mirrored in other mapping datasets is briefly explored. If data elsewhere are consistent with Múlajökull then conclusions drawn in this paper might apply to other sites globally. Of particular interest are whether under-sampling is sufficient to explain the roll-over at other sites (i.e. UK and Sweden), and how under-sampling might distort statistics. So, Section 5.1 employs a correction for under-sampling to verify the existence of roll-overs, whilst Section 5.2 considers the impact of the correction on $\mu_{\rm L}$ as it is commonly calculated [e.g. Ely et al., 2017] and has implications for understanding the rate at which drumlin height equilibrates with ice conditions (Section 6.2.4).

463 5.1 <u>Widespread existence of a roll-over</u>

The widespread existence of a roll-over is tested by the application of a conservative correction for under-sampling [Hillier et al., 2014], in particular the mean curves on Figure 2. As in Section 4.2 the correction is applied to the exponential model to ascertain if it can be bent downwards sufficiently to explain the roll-over. If the roll-over can be replicated by the correction, then the observation is in some doubt. The correction is likely an illustrative 'worst case' for under-detection because the area is hilly and cluttered (i.e. with trees, woods, houses, infrastructure) and the 5 m resolution NEXTmap BritainTM DEM is derived from InSAR observations. Critically, InSAR radar pulses reflect off vegetation, with the returns therefore capturing the top of features such as trees, which must either be statistically removed to estimate a bare earth terrain for mapping [e.g. Sithole and Vosselman, 2004; Clark et al., 2009] or visually compensated for when mapping [e.g. Smith et al., 2006]. This is non-trivial especially, for example, if patches of woodland have greater amplitude and similar spatial scale to drumlins [e.g. Fig 2b of *Hillier and Smith*, 2012]. The correction curve of Hillier *et al.* [2014] is used as it is derived from mapping on synthetic DEMs, and thus the only one available that gives absolute values (i.e. not just relative efficacy between mappers) for under-sampling.

The widespread existence of a roll-over is examined first<u>in UK data</u>. UK drumlins [*Clark et al.*, 2009; *Spagnolo et al.*, 2012] are to a first-order approximated by either a log-Normal or a Gamma distribution for *H*, *W* and *L* [e.g. Fig. 1 of *Hillier et al.*, 2016]. Additionally, for large drumlins above modal size, data are approximated by an exponential tail [*Hillier et al.*, 2013]. Correcting for under-sampling is manifestly inadequate to alter this, namely the correction is not sufficient to explain the roll-over and mapping of few small bedforms in terms of *L* and *W*

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(Figure 8). For *H* the correction (dashed line) to the exponential model (solid line) appears to be
overly pessimistic for drumlins between 5-10 m in relief, which still fit the exponential trend,
and even so cannot explain a relative absence of the smallest forms (Figure 8a). Similar applies
for the flow set of 173 drumlins located at a site near Loch Lomond [*Hillier and Smith*, 2012]
(Figure 9). This is both the site at which the correction was created, and demonstrates that the
observation also applies to flow-set level datasets. All these studies use the NEXTmap Britain[™]
DEM.

Secondly, the roll-over is examined in Swedish data. 20,041 drumlins from Sweden mainly conform to roughly log-Normal size distributions in H, W and L (Figure 10), reaffirming previous statistical analysis [supp. mat. in *Dowling et al.*, 2015]. The roll-over occurs at relatively small sizes in these typically (~94%) rock-cored forms. Mapping was from LiDAR-based DEMs. Again (see Figure 8a), the under-detection correction may be too conservative, but still cannot entirely account for the roll-overs. A factor of x10 (i.e. 10% recovery) equates to 2.3 on the vertical (natural) logarithmic scale so that the fraction of drumlins that would need to be missed to eliminate the roll-overs is considerable.

In short, whether in the UK or Sweden, using InSAR or LiDAR data, for aggregated or individual flow sets, for whichever dimension (i.e. *H*, *W*, *L*), a roll-over exists.

FIGS 8,9,10 HERE

5.2 Distortions to distribution statistics

The extent to which distribution statistics are potentially distorted by under-sampling is examined by using five UK flow sets selected in Ely et al. [2017]. This is not intended as a criticism of this particular data set, which has been extensively and carefully quality controlled [e.g. Clark et al., 2009; Spagnolo et al., 2012]. But, this in itself demonstrates the general applicability of any caution needed. The under-sampling correction is applied to the observations so that any impact on distribution shape can be assessed visually (Figure 11). H is selected for plotting as it is most sensitive to the under-sampling correction. The applicability of a log-Normal distribution [*Elv et al.*, 2017] is to a first order a shown to be valid with or without the correction. The curves, however, are altered with the possibility that the statistics might be notably altered. Visually, the concave-up shape on the below ln(H) of 1 (i.e. ~2.7 m) is not

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1 ว	521	typical of more apparently complete observed distributions (e.g. using LiDAR) or the upper tail						
3	522	of these ones. This is perhaps a qualitative indicator that under-sampling has taken place, and						
4 5	523	\underline{t} he effect of the correction is more pronounced the lower the mode of the flow-set's size						
6 7	524	distribution (see dashed box) <u>.</u>						
8	525							
9 10	526	In terms of quantitative assessment the mean of log-transformed data (i.e. μ_L) is, illustratively,						
11 12	527	the statistic focussed upon. It is relatively commonly used, relates to physically-based statistical						
13 14	528	models, has implications for arguments about rates of equilibrium of drumlins with ice flow						
15	529	(Section 6.2.4), and is intuitively understandable (i.e. location of the distribution). μ_L for L varies						
16 17	530	by ±0.622 (2 σ) (Table 2). This likely reflects a range created by glaciological processes as the						
18 19	531	mean magnitude of the under-sampling correction is 0.035 or ~6% of this. For W the value is 4%,						
20	532	but rises to 56% for H with data varying by ±0.465 (2 σ) and a mean correction of 0.258. Thus,						
22	533	observations of relief might contain materially significant distortions in some cases, perhaps						
23 24	534	contributing to why a trend in $\mu_{\rm L}$ for H predicted by statistical modelling has not been observed						
25 26	535	[<i>Ely et al.</i> , 2017] <u>(Section 6.2.4)</u> .						
27	536							
28 29	537	FIG 11 HERE						
30 31	538	TABLE 2 HERE						
32 33	539							
34	540	<u>6</u> . Discussion						
35 36	541							
37 38	542	<u>6</u> .1 Are small bedforms produced?						
39 40	543							
41	544	The primary purpose of this paper is to rigorously examine the apparent roll-over in drumlin						
42 43	545	size-distributions and the associated scarcity of small drumlins (i.e. < half modal size), which						
44 45	546	might be an important signature of subglacial processes. As a purely empirical descriptor of the						
46 47	547	distribution shape, an exponential distribution fitted above the mode can predict how many						
48	548	small drumlins might naively be expected [Hillier et al., 2013]. Furthermore, a physically-based						
49 50	549	statistical model producing exponential distributions can be conceived [Hillier et al., 2016]. It is						
51 52	550	therefore of interest to determine whether or not under-sampling can explain the existence of						
53	551	the roll-over.						
54 55	552							

the roll-over <u>(Section 4.2)</u> in highly-accurate LiDAR data in an essentially stable, un-vegetated,

 and non-anthropogenically influenced till plain where post-glacial degradation is minimal (Section 3). Mapping in GIS follows best practice, is ground-truthed, and details in the methods of quantification of the morphometrics (e.g. H) are shown to be insufficient to alter this conclusion (Section 4.4). Moreover, even applying a conservative correction for under-sampling during manual mapping is equally insufficient to invalidate the result (Section 4.3). Some doubt might exist as a conservative mapping method was used [Benediktsson et al., 2016; Dowling, 2016], requiring each drumlin to be elevated above its surroundings on all sides. This could miss subtle drumlins that are low relief, but wide and long, because if they are on a gentle larger-scale (i.e. 'regional' [see Hillier and Smith, 2008]) slope then the up-dip face of the drumlin may only be shallower and never actually slope downward. However, this situation is scale-invariant, reflecting only shape and not scale. Thus, such omissions could be argued to reduce the number of small forms found only if small drumlins are also typically flatter; however, there is no evidence that H/W decreases with W either inside or outside the 1992 moraine (i.e. $r^2 < 0.1$). Two other elements of the data presented here also point to a relative scarcity of small drumlins. Firstly, the dominant topographic variations in profiles across the LiDAR DEM viewed at high vertical exaggeration (i.e. x32, inset panels in Figure 5) are at the horizontal scale of drumlins (i.e. ~100 m), without a significant high-amplitude contribution from variations between this scale and ~10 m. Namely, there is no evidence of a progression from the mapped drumlins towards many increasingly small drumlin-like forms. Certainly, it is difficult to see where 10 times the amount that are currently mapped might originate from in order to eliminate the roll-over. Secondly, the mapping [i.e. Benediktsson et al., 2016] suggests that many small drumlins are not going to be revealed by draining the lake-filled interfluves inside the 1992 moraine; this is simply because small drumlins have not been revealed in these locations outside the 1992 moraine. Thus, in summary, at Múlajökull where observational ambiguity is eliminated or minimized, the roll-over is not to be due to i) source data via DEM resolution or quality, ii) 'detectability' or mapper ability in complex (i.e. anthropogenically cluttered or vegetated) landscapes, iii) quantification method to determine morphometrics (e.g. H), or iv) post-glacial degradation. In other words, with preservation and observation accounted for it is possible to clearly state that if small drumlins are created then few survive to pass outside the ice margin. In this sense, they are not 'produced' by the glacier for preservation in the geomorphological record.

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587 <u>I</u>t is also possible to comment on the earl<u>ier</u>-stage, subglacial production of small drumlins.
588 Without sub-ice evidence, it <u>c</u>ould be possible that drumlins originated significantly further up-

stream under the ice and always grew, merged, or were destroyed before being exposed, thereby not ultimately being preserved in the geomorphological record. GPR data preclude this for Múlajökull [*Lamsters et al.*, 2016]. In common with the LiDAR of the forefield (Figure 5), dominant topographic variations in the near-margin basal sub-ice topography are at the scale

592dominant topographic variations in the near-margin basal sub-ice topography are at the scale593of the streamlined ridges that are being interpreted as drumlins (see Figs. 2a & 3 of Lamsters *et*594*al.* [2016]). Namely, there is no evidence of many small drumlins. With respective horizontal595and vertical accuracies of <1 m and 12 m in these GPR data, possible drumlins well within the</td>596roll-over (e.g. *W* ~40 m) would be detectable if they existed. At Múlajökull, at least, it is597therefore not any ice-sediment interaction that occurs during a passage out from under the ice598[e.g. Benediktsson et al., 2016; Lamsters et al., 2016] that eliminates small drumlins and causes599their scarcity.

Finally, Múlajökull [Johnson et al., 2010; Benediktsson et al., 2016; McCracken et al., 2016] provides insights into whether or not many small drumlins ever existed. Everywhere at Múlajökull, including under current ice, small drumlins are scarce. Thus, if any of the three zones (i.e. 'inside', 'outside', currently subglacial) can be argued to reflect the earliest stages of drumlin formation, drumlins must have formed by streamlining pre-existing landforms rather than through progressive growth from small to full-size features. Sedimentology and stratigraphy have been used to create an understanding of the spatial distribution of the cumulative intensity of geomorphic work done by surge-cycles at Múlajökull [Johnson et al., 2010; Benediktsson et al., 2016; McCracken et al., 2016]; specifically, the ice-proximal area inside the 1992 moraine is argued to have experienced more surges, more geomorphological work, and it contains more evolved and elongate (i.e. mean L/W ratio of 3.0) drumlins, than the distal area outside it (mean L/W = 1.9). Also, at Múlajökull the influence of a number of other factors that could affect drumlin morphology (e.g. bedrock, till variation, large-scale topography) is likely minimal. Thus, since the Múlajökull drumlins have demonstrably elongated in surges, and yet the distal ones have not elongated much (i.e. L/W < 2.0), it is possible to argue that the distal drumlins are comparatively geomorphologically 'immature' and represent an early stage of drumlin formation. As such, observing few small drumlins in the till plain outside the 1992 moraine strongly implies that few were produced or existed in the earlier stages of drumlin genesis. A first explanation for the scarcity of small drumlins in apparently immature zones (e.g. elongation ratio ≤ 2.0), at least at sites similar to Múlajökull (i.e. lobes with near-margin drumlin genesis), is that drumlins form by streamlining pre-existing landforms (e.g. moraines, debris fans) rather than through progressive growth from small to full-size

features. A second explanation is that small bedforms grew and/or merged rapidly before significant streamlining had occurred. In terms of all H, L and W small drumlins are scarce even in the apparently immature area outside the 1992 moraine [Benediktsson et al., 2016], which has low elongation ratios (i.e. <2.0), so if they once existed they must have disappeared by this stage. A third, and our preferred, explanation is that drumlins at Múlajökull initiate as relatively broad and shallow features that increase in amplitude notably faster than they elongate, and grow and/or merge rapidly before significant streamlining has occurred. In addition to the other constraints (e.g. such as mean sizes $\mu_{\rm L}$ in Section 6.2.1), this is supported by an interpretation of the spread of size-frequency observations (i.e. $\sigma_{\rm L}$) in the context of statistical modelling that give insights into rates of change (see Section 6.2.3). In the later two explanations, regularity could emerge through the aspects of the process of growth (e.g. stochasticity) rather than reflecting initial conditions [see Hillier et al., 2016]. In the first explanation, pre-existing features with some regularity in spacing, are required to conform with this tendency in drumlins [Clark et al., 2018].

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Away from Múlajökull, the complication exists that drumlin fields are produced substantially away from an ice margin. Thus, the possibility exists for sub-ice modification or destruction of any small drumlins produced before they emerge. Additional sub-ice evidence (e.g. GPR) is needed to constrain this possibility further. Seismic [King et al., 2007] and radar [King et al., 2009] data collected in Antarctica are of a different character and lower horizontal resolution (~50 m) than LiDAR, and of mega-scale glacial lineations (MSGL), but it is interesting that they are reported as visually indistinguishable from relict bedforms of the Dubawnt palaeo-ice stream bed captured in Landsat images. The Dubawnt flow set lacks the many small bedforms expected of the exponential extrapolation [Hillier et al., 2013; Stokes et al., 2013a], and combining this with the Antarctic geophysics gives a first tentative indicator of potential scarcity under ice streams. A final possibility, that is difficult to constrain, is that Múlajökull's foreland may simply represent an atypical phase of drumlin field evolution where small drumlins are under-represented.

Observational certainty is higher for forms that are currently exposed. A compilation of various datasets from the UK and Sweden [Clark et al., 2009; Hillier and Smith, 2012; Spagnolo et al., 2012; Dowling et al., 2015; Ely et al., 2017] in Section 5 demonstrates that the roll-over exists for data from varied locations, with varied data sources, and for both aggregated data and that at the level of individual flow sets, which potentially represent glaciological conditions in a

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single place and time. Predominantly, this is valid even with a correction for under-detection applied (Section 5). Thus, behaviours similar to that at Múlajökull (i.e. whatever leads to few small forms) may be more widespread than just that site. The roll-over is least apparent in rock-cored drumlins recorded in LiDAR data [i.e. Dowling et al., 2015] where, in contrast to other data, a log-Normal distribution does not well explain the very largest forms and a Gamma distribution is visibly less adequate (Figure 10). It therefore remains entirely possible that the balance of physical processes that dictate bedform sizes changes along a spectrum from hard-cored to soft-cored features, with small rock-cored features being easier to produce.

<u>6.2</u> What can size-frequency observations say about drumlin evolution?

With the robustness of the size-frequency data at Múlajökull established (Sections 4 & 5), they can be taken as a firm constraint upon drumlin formation models (i.e. statistical, conceptual, or numerical ice flow) intended to apply at this site. It is open to debate how representative the Múlajökull site is, but it must be incorporated for a theory to have general applicability, and so the implications of the statistical size analysis are discussed below with this taken as read.

The current conceptual model for Múlajökull [Benediktsson et al., 2016; McCracken et al., 2016] is considered first (Section 6.2.1). After this only the 132 currently exposed drumlins are considered as, at present, they are the observations that most reliably isolate the evolution of bedforms through time with other conditions held constant (see Section 2). Early, protodrumlin morphology is noted in Section 6.2.2, with implications for the initiation of any numerical or statistical model. Then, the applicability or otherwise of statistical models to the Múlajökull site is evaluated, constraining which remain tenable. Finally, rates at which the dimensions (i.e. H, W, or L) equilibrate is considered, contributing to our overview of how drumlin formation functions.

6.2.1 Múlajökull conceptual model

The current conceptual model of [Johnson et al., 2010; Benediktsson et al., 2016; McCracken et al., 2016] (see Section 2) is based on sedimentary and stratigraphic observations, and the reported mean changes in H, W and L at Múlajökull. Thus, it cannot be tested by these changes, but the increase in L with exposure to more surges (i.e. inside the 1992 moraine), decrease in Wand probable invariance in H are all verified statistically (Table 1, Section 4). Namely, the model

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691 <u>is not based on a falsely confident misinterpretation of summary statistics from small samples</u>
 692 <u>of data in the foreland.</u>

 The five drumlins measured under Múlajökull by Lamsters et al. [2016] tentatively suggest that the trends observed in the foreland do not continue beneath the glacier; they are higher and wider rather than narrow and of similar height as simple extrapolation would suggest. Without detailed explanation this was attributed to either i) more till layers, ii) more sediment due to unspecified ice-stress differences, or iii) variation in till composition (e.g. stiffer till more resistant to erosion). A smoother transition appears to exist between the swales and crests of the current subglacial drumlins than on the foreland. Without yet being subject to being interspersed by proglacial lakes and their associated sedimentation, these subglacial drumlins lack clear breaks in slope at their margins, and thus the former glacier bed is likely not directly comparable to its foreland [e.g. Finlayson, 2013]. Some metrics (e.g. inter-crest spacing) appear little affected and similar to the exposed drumlins, whilst H in particular is more sensitive. It is clear, however, that the model of Múlajökull cannot ultimately only be based on data from the foreland.

<u>6.2.2</u> Initial conditions for <u>numerical and statistical</u> modelling

Observations at Múlajökull imply that drumlins may not initiate as perturbations that are very small in all dimensions (i.e. H, L, W), a simplifying assumption commonly used in both numerical ice-flow [e.g. Hindmarsh, 1998; Dunlop et al., 2008; Chapwanya et al., 2011] and statistical [Fowler et al., 2013; Hillier et al., 2016] modelling. The current model for Múlajökull [Johnson et al., 2010; Benediktsson et al., 2016; McCracken et al., 2016; Iverson et al., 2017] postulates that drumlins initiate as wide, rounded and relatively low amplitude topographic features (i.e. <u>H</u> is small but L and W are not), and the inferences from the size-frequency observations at Múlajökull in the context of the Stochastic Instability statistical model are consistent with this (see Section 6.2.3). Exploring fully the implications of this initial condition used in models is outside of the scope of this paper, so it is simply noted that future modelling should consider the sensitivity of outputs to the initial size distribution selected for the modelling [e.g. see Hillier et al., 2016].

723 <u>6</u>.2.<u>3</u> Statistical models

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As a step to bridging the gap between geomorphological form and process Hillier et al. [2013] proposed a conceptual model to explain the size-distributions of subglacial bedforms in terms of stochastic ice-sediment-water interaction. This has led to a variety of physically-based statistical models being developed to formalize the postulated stochastic behaviour [Fowler et al., 2013; Hillier et al., 2016]. The statistical models include various elements (e.g. initial size distribution, growth rate law) and predict size-frequency distributions and how they evolve through time as the drumlins evolve. In making specific predictions about observable quantities (e.g. μ_L), they are testable and falsifiable.

A statistical model based on waiting time (i.e. Poisson) randomness and a single episode of drumlin building, that might be a surge, can produce an exponential size-frequency distribution. Illustratively, this is model M8 of Hillier et al. [2016]. However, a securely evidenced roll-over at Múlajökull (Section 4.2) and more widely (Section 5) now firmly precludes any statistical model that produces an exponential size-frequency distribution from being a viable model. Two statistical models that include randomness through time in drumlin growth, and are based on glaciologically plausible physical conditions, can explain size-frequency distributions with a roll-over, but these are difficult to distinguish in aggregated UK data [Hillier et al., 2016]. A powerful aspect of the Múlajökull site is that it contains a progression from less evolved drumlins outside the 1992 moraine to more evolved ones inside, effectively two snapshots of drumlin growth at two different times. This constraint allows tests of predictions of how subglacial bedforms (e.g. drumlins) progressively evolve with time, which gives more potential to distinguish between models.

The first physically-based statistical model that can explain a roll-over is the Waiting Time (WT) model [Hillier et al., 2016]. The WT model is based on Poisson randomness and creates a Gamma distribution with two parameters (i.e. α , β). β is the rate at which conditions in the ice-sediment-water switch between those suitable for growth and those that cause bedforms to shrink, and is expected to remain constant through time in the WT model as constructed. No change in β with time for H and W at Múlajökull is therefore in agreement with the WT model, but the statistically significant decrease for L is difficult to reconcile with the model. α reflects the number of growth episodes (e.g. surges) experienced on average by the bedforms, and so is expected to grow with time. Thus, a tendency to decrease in H and W even through not statistically significant, especially combined with the statistically significant decrease in α for L, produces an observation that is inconsistent with the WT model. The WT model as currently

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constructed is therefore falsified, although it is worth emphasizing that as with mathematical
 models of drumlin formation [e.g. Hindmarsh, 1998; Chapwanya et al., 2011; Hooke and
 Medford, 2013; Iverson et al., 2017] variants that remedy this might be constructible.

The second physically-based statistical model that can explain a roll-over is the Stochastic Instability (SI) model created by Fowler et al. [2013] and re-formulated and generalized by <u>Hillier et al. [2016]. The SI model</u> creates log-Normal distributions with two parameters (i.e. μ_L , σ_L). μ_L has already been interpreted at Múlajökull in terms of the state of the drumlin sizes at two times (Section 6.2.1), but σ_L can offer additional information on the rate of changes at the two times. σ_L for H is greater than that for W or L implying that growth rate (k) is fastest in this dimension both inside and outside the 1992 moraine [see Eq. 27 of Hillier et al., 2016], in agreement with aggregated UK data [Hillier et al., 2016]. Outside the 1992 moraine, where fewest surges have sculpted the morphology, σ_L values imply k_W exceeds k_L . At face value, this implies that drumlins outside the 1992 moraine are getting less elongate with time. Alternatively, it can be interpreted as being consistent with W values not initially starting small as assumed in the SI model, within which the only way to become wide is to grow to be wide. This second interpretations is much easier to reconcile with the Múlajökull site (see Section 2). Inside the 1992 moraine σ_L values imply k_L exceeds k_W as for the UK data [i.e. of Clark et al., 2009], indicating that drumlins have elongated with time. Observations in the two zones can be reconciled if drumlins elongate with time, but with streamlining taking a little time to become dominant as the signal of the initial conditions (i.e. broad gentle proto-drumlins) is progressively over-printed. This is entirely consistent with suggestions in Hillier et al. [2016] that different dimensions might behave differently (i.e. L continuing to grow after W is restricted). Overall, the observations do not falsify the SI model, indeed the most logical interpretation of the size-frequency data places it into agreement with initial conditions recently postulated in mathematical model for drumlin formation at Múlajökull [e.g. Johnson et al., 2010; Iverson et al., 2017]. Thus, it is clear that statistical models will have greater explanatory power if used in conjunction with site-specific conceptual or mathematical models.

<u>6</u>.2.<u>4</u> Equilibration with flow conditions

Ely *et al.* [2017] speculatively interpret a lack of trend in μ_L and σ_L for *H* in UK flow sets as rapid stabilisation, consistent with inferences by Hillier *et al.* [2016] using aggregated UK size data that *H* grows and evolves relatively more rapidly that *W* or *L*. The other likely explanation for an

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absence of a trend in the data of Ely et al. [2017], which cannot yet be excluded (Section 5.2), is observational uncertainty in H (e.g. due to under-detection or the measurement technique used); i.e. a trend might still be present, just masked by noise. However, the lack of a distinguishable change in H between two areas either side of the 1992 moraine at Múlajökull that have experienced similar glacial conditions but a different number of surges adds weight from a better constrained site to the view that H stabilises rapidly. Specifically, stabilisation here is refined to mean that H has finished increasing or decreasing, does this rapidly with respect to L or W, and perhaps implies equilibrating with ice-flow conditions. In this context rapid must mean short compared to the ~400-800 year LIA time frame available at Múlajökull indicated by formation of the Arnarfellsmúlar terminal moraine. This adjustment might even be on the decadal timescale [e.g. Hillier et al., 2016], based on geophysical evidence [e.g. Smith et al., 2007], sediment flux [Rose, 1989] and geometrical arguments [Goldstein, 1994; Dowling et al., 2016]. Stabilising H in the absence of a sharp spike in observational frequency at a postulated capping height requires any upper limit on dimensions to be 'fuzzy' or probabilistic [*Hillier et al.*, 2016].

7. Conclusions

From statistical analysis of 143 newly emergent drumlins, recently created by the Múlajökull glacier, in conjunction with on the order of 100,000 drumlins mapped in the UK and Sweden, the following main conclusions can be drawn.

Few small drumlins are produced in the active Múlajökull drumlin field.

Our preferred explanation for the scarcity of small drumlins at Múlajökull is that drumlins form stochastically, during surge cycles, by wide and gentle pre-existing undulations rapidly increasing in amplitude before significant streamlining occurs. The scarcity of small drumlins in the less evolved (i.e. 'immature') zone outside the 1992 moraine requires the rapidity, whilst size-frequency observations (i.e. μ_L , σ_L) in the context of statistical modelling give insights into rates of growth and imply the protodrumlin morphology.

In the UK and Sweden, and in a variety of data types, size-frequency distributions have a 'roll-over' similar to that at Múlajökull, providing wider_evidence that few small drumlins exist in previously glaciated landscapes. Thus, behaviours similar to that at Múlajökull may be more widespread than just that site.

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It is also interesting to note that the first-order properties (i.e. approximately log-Normal shape) of size-frequency distributions is likely not altered substantially by under-sampling, although most care is needed for relief (i.e. height). Size-frequency distributions for H remain most sensitive to under-detection, and might still have materially significant distortions in some cases, which should be accounted for when interpreting data or derived statistics. Finally, it seems clear that statistical models are useful companions to their numerical ice-flow counterparts as tools to assist our understanding of ice-base processes, especially if placed into a site-specific context. For example, a mathematical model focused on physical behaviours at Múlajökull [e.g. Johnson et al., 2010; Iverson et al., 2017] might be blended with statistical modelling (i.e. Stochastic Instability model) [Hillier et al., 2016]. With a statistical model ground-truthed at a site, the modelling could be used to extrapolate and thereby be tested for consistency with observations across Earth. In particular, observations at Múlajökull add weight to a model previously posited for testing by *Hillier et al.* [2016] in which H evolves relatively rapidly [Hillier et al., 2016] to be at equilibrium with ice-sediment-water conditions [Ely et al., 2017], W changes more slowly constrained geometrically by interactions with neighbouring bedforms [e.g. Hillier et al., 2013, 2016; Clark et al., 2018], but L is free to grow. Thus, statistically enhanced modelling could feed into the longstanding debate on a subglacial bedform continuum [e.g. Aario, 1977; Rose, 1987] and be a step towards using drumlin fields as proxies for the critical parameters used in ice sheet reconstructions and modelling.

Acknowledgements

We thank the organisers of the 'Beauty of Drumlins' symposium for bringing the authors
together in Lund, Sweden, May 2017, in honour of Professor Per Möller upon his retirement.
We are grateful to Chris Stokes and an anonymous reviewer for their insightful comments,
which improved the manuscript.

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Table 1: Parameters of the drumlin data from Múlajökull. Shading indicates a statistical significance of differences; light shading *P* < 0.1, dark shading *P* < 0.05. Bold numbers indicate if skew is different from 0 (P < 0.05). Uncertainties are 1σ , and tests 1-tailed. Note that all parameters are estimated from the underlying data, and binning for visualization (i.e. Figure 7) is solely for that purpose.

Distribution	⁴ Parameter		² H	L		W			
		Inside	Outside	Inside		Outside	Inside		Outside
Normal	³ μ (m)	7.6	≈ 7.4	219.5	>	168.9	80.5	<	93.8
		±0.3	±0.3	±9.4		±5.8	±2.7		±4.0
	skew	0.258	0.193	0.861		0.764	0.228	<	1.011
		±0.153	±0.214	±0.379		±0.248	±0.250		±0.397
Log-Normal	μ	1.963	1.940	5.373	>	5.099	4.342	<	4.495
		±0.045	±0.048	±0.041		±0.033	±0.036		±0.041
	σ_{L}	0.393	0.354	0.362	>	0.245	0.315		0.303
		±0.032	±0.043	±0.029		±0.021	±0.027		±0.031
Gamma	α	7.270	9.029	8.074	<	16.908	11.046		11.201
		±1.146	±1.691	±1.265		±3.186	±1.753		±2.103
	β	0.952	1.226	0.035	<	0.100	0.137		0.119
	-	±0.155	±0.236	±0.006		±0.019	±0.022		±0.023
	¹ φ (i.e. mode)	6.3	6.4	210.5		161.6	72.3		82.5
Exponential	3λ	0.359	0.417	0.013	<	0.026	0.046	>	0.032
		±0.050	±0.072	±0.002		±0.005	±0.007		±0.006
Non-parametric	⁵median	7.5	7.0	220	>	164	79	<	85

¹ estimated from the Gamma distribution as Hillier *et al.* [2013], but using maximum likelihood estimates. Derived from α and β so significance of any differences not estimated. ² 1992 and 2008 moraines removed if *H* increased by >2 m

³ Significance calculated using Welsh t-test, 2-tailed; $1/\lambda$ is a mean [e.g. see *Hillier et al.*, 2013].

⁴ Unless otherwise stated, significance by non-parametric bootstrapping; resampling is with replacement within sub-sets, n = 10,000.

⁵ Two-sample Wilcoxon test; strictly, a non-parametric test of difference in distribution location,

not difference in medians, but it is a useful and relevant indicator.

1118 Table 2: Means of the logarithms of size data (i.e. μ_L) for various global data, both before and

1119 after a conservative correction for the completeness of mapping is applied i.e. Fig. 1c of [Hillier

1120 et al., 2014].

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1126	Figure	Captions
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Figure 1: The typical shape of size-frequency distributions for mapped drumlins. These probability density functions are similar to normalized histograms, and are plotted on semi-log axes for graphical purposes (i.e. an exponential distribution plots as a straight line). a) Size-frequency data for two studies, as black dots [Clark et al., 2009] and a grey line [Hillier and Smith, 2012], illustrated with length. Selected statistical distributions are fitted to them: exponential distribution (solid blue line); gamma distribution (dashed line) [Hillier et al., 2016]; log-Normal (dotted lines)[Fowler et al., 2013]. b) The typical distribution shape [Hillier et al., 2013], with possible explanations for it annotated in bold text next to the part of the distribution they may impact.

> Figure 2: Recovery rate (i.e. 'completeness') as a function of size, with the use of realistic synthetic DEMs allowing absolute values for recovery to be determined [Hillier et al., 2014]. Solid black line is for height, H, and grey lines are for width W (solid) and length L (dashed). Circles are means, shown with their standard errors across 10 synthetic DEMs. Dashed black line is for medians for H. H, W, and L have bin widths of 2.5, 25 and 100 m, respectively. At the upper end, bins with two or fewer input data are omitted, giving maxima of 20, 275 and 800 m, respectively. All data are plotted centrally within bins.

Figure 3: a) Location of Múlajökull (white square) on the southern edge of the Hofsjökull ice cap. b) Overview photograph, July 2011, with view from the SE. Photo courtesy of Sverrir A. Jónsson. c) Múlajökull glacier and surrounding area, adapted from Benediktsson et al. [2016], who also detail the data and mapping methods. Background map is the 2013 LiDAR hillshade in a mosaic with 2014 orthophotos. Drumlins (red with white outlines) are all within limits of the 2013 LiDAR data coverage. P1 and P2 are terrain profiles shown in Figure 5. The LIA, 1992, 2008 and 2013 ice limits are labelled. Dashed white line on ice surface indicates the approximate edge of the overdeepening beneath Múlajökull and the up-glacier limit of the drumlin field [Lamsters et al., 2016]. UTM (zone 27N) coordinates used. d) A drumlin emerging from the central margin of Múlajökull in 2014. Ice flow is towards the viewer.

Figure 4: a) Extract from a 2014 orthophoto showing the sparsely to non-vegetated area inside

the LIA moraine and the contrast to continuous vegetation beyond it. b) Extract from a 2015 URL: http://mc.manuscriptcentral.com/sqff Email: Christian.Skovsted@nrm.se

high-resolution orthophoto recorded with an unmanned air vehicle showing exposed and nonvegetated drumlins in front of the ice margin in 2015 (photo courtesy of Jez Everest, British
Geological Survey). c) View of the distal slope of the LIA terminal moraine, exemplifying the
continuous but low vegetation on and beyond the moraine. Ice flow was from left to right. The
location of the photograph is indicated with a black asterisk on a).

Figure 5: Profiles illustrating DEM data quality of the 2008 (grey line) and 2013 (black line) LiDAR surveys, and stability through time of the till plain. Spatial extent is limited to inside the 1992 moraine as the 2008 data only extend that far. Thin vertical lines are the limits of the numerical comparison, dashed lines the trends fitted by ordinary least squares to the 2013 data, and grey horizontal bars indicate where 2013 heights are 'high' (i.e. above the trend).

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> Figure 6: Elevation change between 2008 and 2013 LiDAR DEMs where they overlap (i.e. inside the 1992 moraine). <u>All changes shown in colour are decreases in elevation, whilst greys within</u> the polygon of coincident LiDAR are increases. Black and blue solid lines show the 2008 and 2013 ice margins, respectively. Background map is the 2013 LiDAR hillshade. Profiles in Figure 5 are P1 and P2. Coordinate system is UTM zone 27.

Figure 7: Semi-log size-frequency plots for *H*, *W*, *L* for inside (a-c) and outside (d-f) the 1992 moraine at Múlajökull with exponential (grey line), Gamma (blue dashed line) and log-Normal (pale blue dotted line) distributions fitted. Data (black dots) are from <u>Benediktsson et al.</u> [2016] (*H*, *W*, *L* triplets given in Supplementary Material), binned to illustrate the empirical density function, and distribution parameters are in Table 1. Extent to which a conservative correction for detectability and mapping ability (i.e. Figure 2 or Fig. 1c of [Hillier *et al.*, 2014]) can influence the exponential model is shown as a grey dashed line.

Figure 8: Potential for incomplete mapping to explain the roll-over in a) height, b) length, and c) width for aggregated UK data [Clark et al., 2009; Spagnolo et al., 2012]. Data are black dots. Plots are semi-log and of count density, with an exponential model fitted to drumlins larger than the mode as Hillier et al. [2013] (grey line). Exponent is λ . Extent to which a conservative correction (i.e. Figure 2 or Fig. 1c of [Hillier et al., 2014]) can influence this model is show (grey dashed line); if dashed line descends below the data at smaller sizes this indicates a potential magnitude exceeding that of the roll-over.

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1 2	1194	Figure 9: Potential for incomplete of mapping to explain the roll-over in a) height, b) length, and
3	1195	c) width for drumlins near Loch Lomond [Hillier and Smith, 2012].
4 5	1196	
6	1197	Figure 10: Potential for incomplete of mapping to explain the roll-over in a) height, b) length,
8	1198	and c) width for drumlins Swedish drumlins; extended dataset based on Dowling et al. [2015].
9 10	1199	d), e) and f) expand section of these plots around the roll-over. Exponential (grey line), Gamma
11 12	1200	(blue dashed line) and log-Normal (pale blue dotted line) distributions fitted. Possible
13	1201	correction for detectability is as in Figure 7.
14 15	1202	
16 17	1203	Figure 11: Probability distributions for log-transformed heights (H) for selected UK flow sets
18	1204	from Elv et al. [2017] (Fig. 7f), in which a Gaussian shape illustrates a log-Normal distribution, a)
19 20	1205	is uncorrected observations, whilst b) has had the under-sampling correction applied.
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1 2 3 Drumlins INSIDE the 1992 moraine 4 5 Height (H), Width (W) and Length (L) triplets for the 132 fully exposed drumlins as mapped 6 and estimated by Benediktsson et al (2016). Each HWL triplet is on a row. 77 7 drumlins. 8 All measurements are in metres. 9 10 Benediktsson et al [2016] Boreas, 45, 567-583. 11 12 Column 1 = Height(m)Column 2 = Width (m)13 Column 3 = Length(m)14 15 8.200 59.000 169.000 16 4.600 28.000 117.000 17 2.600 41.000 74.000 18 9.200 86.000 166.000 19 4.300 47.000 145.000 20 8.100 93.000 133.000 21 7.200 80.000 250.000 22 6.100 78.000 165.000 6.300 74.000 220.000 23 9.200 99.000 275.000 24 7.000 79.000 278.000 25 4.900 78.000 164.000 26 7.700 95.000 320.000 27 5.900 68.000 239.000 28 11.900 118.000 236.000 29 9.000 75.000 305.000 30 7.800 67.000 200.000 12.100 75.000 254.000 31 8.900 46.000 349.000 32 4.100 63.000 133.000 33 5.700 74.000 236.000 34 12.400 82.000 267.000 35 3.000 40.000 162.000 36 6.900 67.000 216.000 37 8.500 80.000 273.000 38 5.600 67.000 184.000 39 8.700 78.000 262.000 7.800 76.000 335.000 40 11.700 107.000 368.000 41 7.000 62.000 220.000 42 8.100 77.000 251.000 43 8.000 96.000 384.000 44 7.700 62.000 192.000 45 7.700 86.000 244.000 46 4.100 51.000 179.000 47 6.800 83.000 206.000 7.500 102.000 225.000 48 10.300 103.000 354.000 49 8.900 95.000 221.000 50 11.500 115.000 361.000 51 2.500 47.000 144.000 52 7.200 85.000 259.000 53 3.800 57.000 116.000 54 11.400 95.000 344.000 55 6.600 101.000 276.000 5.300 63.000 185.000 56

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