Prescriptive variability of drugs by general practitioners

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1 Abstract

2 Prescription drug spending is growing faster than any other sector of healthcare. 3 However, very little is known about patterns of prescribing and cost of prescribing 4 between general practices. In this study, we examined variation in prescription rates 5 and prescription costs through time for 55 GP surgeries in Northern Ireland Western 6 Health and Social Care Trust. Temporal changes in variability of prescribing rates and 7 costs were assessed using the Mann-Kendall test. Outlier practices contributing to 8 between practice variation in prescribing rates were identified with the interguartile 9 range outlier detection method. The relationship between rates and cost of prescribing 10 was explored with Spearman's statistics. The differences in variability and mean 11 number of prescribing rates associated with the practice setting and socioeconomic 12 deprivation were tested using t-test and F-test, respectively. The largest between-13 practice difference in prescribing rates was observed for Apr-Jun 2015, with the 14 number of prescriptions ranging from 3.34 to 8.36 per patient. We showed that practices with outlier prescribing rates greatly contributed to between-practice 15 16 variability. The largest difference in prescribing costs was reported for Apr-Jun 2014, with the prescription cost per patient ranging from £26.4 to £64.5. In addition, the 17 18 temporal changes in variability of prescribing rates and costs were shown to undergo 19 an upward trend. We demonstrated that practice setting and socio-economic 20 deprivation accounted for some of the between-practice variation in prescribing. Rural 21 practices had higher between practice variability than urban practices at all time points. 22 Practices situated in more deprived areas had higher prescribing rates but lower variability than those located in less deprived areas. Further analysis is recommended 23 24 to assess if variation in prescribing can be explained by demographic characteristics of patient population and practice features. Identification of other factors contributing 25

to prescribing variability can help us better address potential inappropriateness ofprescribing.

28 Introduction

In recent years, NHS spending on drugs has substantially risen, from £13.0 billion in 29 2010/11 to £16.8 billion in 2015/16 [1]. Most of the expenditure on prescribed 30 31 medicines is incurred in primary care and closely related to the steadily growing 32 workload of general practitioners (GPs) [1]. In England, patient consultations with GPs increased by 16% in the period 2007-14 [2] whereas in Northern Ireland, the GPs 33 34 workload grew by 22% over the same period [3]. In addition, there has been an 35 approximately 60% increase in prescription items dispensed from 2005 to 2014 in 36 Northern Ireland [3] and a corresponding 50.4% rise in the number of prescriptions dispensed in England [4]. 37

The National Audit Office report found that substantial savings for the NHS could be 38 achieved by improving the overall quality and cost-effectiveness of prescribing [5]. 39 Accordingly, a lot of interest has been focused on variation in prescribing practice as 40 a potential source to save money [5,6]. Despite a wealth of literature on prescribing 41 42 patterns [5,7,8,9], there is a lack of full understanding of factors that contribute to between-practice differences in prescribing. Among key influences upon prescribing 43 44 variation, the demographic and socio-economic characteristics of patient population 45 (e.g. age, ethnicity, deprivation) are most often acknowledged by researchers [10,11]. GP practices with a greater proportion of people in older age groups were more likely 46 to prescribe minor tranquilisers [10], sex hormones, anticoagulants and protamine, 47 48 and treatments for glaucoma [12]. Significant differences in prescribing were also

49 associated with the level of deprivation [13]. Several studies have shown that extent 50 of local deprivation influences antidepressant and lipid-lowering medication prescribing [14,15,16]. On the other hand, lower volume of prescribing was observed 51 52 in practices with higher proportions of patients from ethnic minority populations [17]. Practice features were also among factors contributing to the variation in prescribing 53 54 behaviour. Examples of that include higher prescription rates issued by practices 55 located in urban areas with a greater proportion of female GPs [18]. Lower prescribing was found for single-handed practices, practices in rural areas, with a higher average 56 57 age of general practitioners, and with GPs born outside the UK [15,19,20].

58 Differences in characteristics of GP practices or a patient population do not always 59 explain GPs prescribing behaviour. In many cases, the variability in prescribing rates is associated with inefficient or inappropriate prescribing [5,21]. It has been estimated 60 that the prescription costs could be reduced by as much as £1bn if unwarranted 61 62 variations in prescribing levels were eliminated and the drugs were prescribed with the same standard [21]. Better efficiency and appropriateness in prescribing practice could 63 64 be achieved by addressing the over- or under-utilisation of drugs. It was shown that 65 prescribed medications are often taken for long periods beyond the point when they are needed and around 30% of drugs are abandoned by patients [22,23,24,25]. Major 66 NHS savings could also be generated by using treatments that are most cost-effective. 67 Moon et al. [26] showed that a large number of GPs are still prescribing brand name 68 medications, even though the cheaper, equally safe and effective alternatives are 69 70 available.

The aim of this study was to investigate temporal changes in rates and costs of prescribing as well as between-practice variation in prescribing. In addition, we examined if prescribing rates of GPs were related to the practice setting and
 socioeconomic deprivation.

75 Methods

76 Data and pre-processing

We analysed the number and actual cost of prescription items issued by 55 general practices within the Northern Ireland Western Health and Social Care Trust (WHSCT) during twelve consecutive periods of 3 months, starting from Apr 2013 to Mar 2016. The actual cost of prescriptions was defined as the estimated cost to the NHS calculated by subtracting the discount per item from the gross cost which is the basic price of a drug.

The GP prescribing data was obtained from the Business Services Organisation's (BSO) prescribing and dispensing information systems [27]. It includes prescribing for all GPs and other non-medical prescribers who are attached to GP practices i.e. nurses, pharmacists, optometrists, chiropodists, and radiographers. To allow temporal comparison of prescribing data, the number of drug prescriptions and their total cost calculated for each general practice was adjusted for the total number of patients in each practice and expressed as prescriptions/cost (£) per patient.

Given data from the Census Office of the Northern Ireland Statistics and Research
Agency [28], a practice was designated as urban if its postal address was situated in
a settlement of more than 10,000 residents. Under this definition, 31 practices were
categorised as urban and 24 as rural.

In addition, practices were categorised based on the Northern Ireland Multiple
Deprivation Measure (NIMDM) at the level of Super Output Area (SOA) [29]. The

96 NIMDM consists of seven domains i.e. Income; Employment; Health, Deprivation, and 97 Disability; Education, Skills and Training; Proximity to Services; Living Environment; and Crime and Disorder. On this overall measure, the SOA with a NIMDM rank of 1 is 98 99 considered the most deprived, and 890 the least deprived. Accordingly, a practice situated in SOA with the NIMDM rank larger than 445 was designated as 'located in a 100 101 less deprived area' while a practice situated in SOA with a NIMDM rank smaller than 445 was designated as 'located in a more deprived area'. Under this definition, we 102 103 identified 11 practices 'located in less deprived areas' and 44 practices 'located in 104 more deprived areas'.

105 Statistical analysis

106 The variation in the number and cost of prescriptions per patient was assessed by 107 calculating the variance (σ^2) for each of the 12 considered time points [30]. In addition, 108 we analysed changes in mean (μ) and range of the rate and cost of prescriptions.

The outlier GP practices were identified for all time points using the interquartile range (IQR) method for outlier detection [31]. Accordingly, a practice with the prescribing rate that fell outside either 1.5 times the IQR below the first quartile or 1.5 times the IQR above the third quartile, was considered to be an 'outlier'. We however acknowledge that a statistical outlier in terms of prescribing rate is not necessarily an example of inappropriate practice.

The differences in the mean number of prescribing rates, for the rural and urban practices as well as practices located in areas of different levels of socioeconomic deprivation, were assessed using an unpaired t-test [31]. The equality of variances of prescribing rates for above-mentioned practice categories was evaluated using F-test [30]. The normality of prescribing data was confirmed with Shapiro-Wilks test [32]. To determine if temporal changes in variability of rates and costs of prescribing underwent a statistically significant upward or downward trend over the study period, we used the Mann–Kendall test which has been commonly employed to detect trends in series of data [33,34].

The relationship between rates and cost of prescribing was explored with Spearman's rank correlation (*rho*) [35]. We chose the Spearman correlation measure due to it insensitivity to individual contribution of outliers. The strength of correlation was defined as very weak for |rho| = 0.2 to 0.39, moderate for |rho| = 0.4 to 0.59, strong for |rho| = 0.6 to 0.79, and very strong for |rho| = 0.8 to 1 [35].

129 **Results**

The total number of patients registered at 55 general practices providing services throughout 2013–16 increased from 318,057 in 2013-14 to 326,429 in 2015-16. Over this time, the total actual prescription cost continued to rise from £58,669,971 in 2013-14 to £63,803,168 in 2015-16.

Fig 1 shows the magnitude and temporal changes in variability of the number of 134 prescriptions per patient. We observed large differences in drug prescribing rates 135 136 among individual practices. The largest between-practice difference in prescribing rates was observed for the quarter of Apr-Jun 2015, with the number of prescriptions 137 138 ranging from 3.34 to 8.36 per patient. During this period, the prescription rate for the 139 practice with the largest number of prescriptions per patient was ~ 60% higher than the average prescribing rate for all the practices ($\mu = 5.20, 95\%$ CI = [4.96.5.44] 140 prescriptions per patient). The smallest between-practice difference in prescribing 141 142 rates was observed in the period Apr-Jun 2013, with the number of prescriptions 143 ranging from 3.21 to 7.60 per patient. At that time, the practice with the highest 144 prescribing rate issued ~ 49% more prescriptions per patient compared to the average prescribing rate of $\mu = 5.11$, 95%CI = [4.89, 5.33]. The high inter-practice variability in 145 146 drug prescribing behaviour was caused by: 1.8% (Oct-Dec 2013), 3.6% (Apr-Jun 2013, Oct 2014 - Mar 2016), 5.5% (Jan-Sep 2014), and 7.3% (Jul-Sep 2013) of GP 147 148 practices with outlier prescribing rates. By eliminating the effect of these outliers (i.e. practices with higher or lower prescribing rates than the calculated outlier cut-off 149 150 values), we were able to reduce the between-practice variability in prescribing rates 151 from 21% (σ^2 reduced from 0.71 to 0.59 in Oct-Dec 2013) up to 70% (σ^2 reduced from 152 0.67 to 0.39 in Jul-Sep 2013) (S1 Table). It is worth highlighting that despite varying 153 number of outliers identified in each quarterly period, they were mostly the same 154 practices: one practice (with substantially higher prescribing rate than outlier cut-off 155 values) was identified as an 'outlier' throughout the studied period while two other 156 practices (one with higher and the other with lower prescribing rate than outlier cut-off 157 values) were labelled as 'outliers' at 11 and 4 considered time periods respectively.

Fig 1. Temporal variability in the standardized number of prescriptions. Each data point (dot): a single practice. Solid, horizontal line inside the box: median of data. Green diamond: mean. Lower and upper "hinges" of the boxplots: 1st and 3rd quartiles, respectively. Red, green, and blue lines: trend lines for maximum, average, and minimum values of prescription rates respectively. Lower and upper extremes of whiskers: interval boundaries of the non-outliers (black dots). Data outside interval (red dots): outliers.

165 Temporal variability in the actual cost of prescribed medications per patient is shown
 166 in Figure 2. The largest between-practice difference in prescribing costs was observed

167 for the quarter of Apr-Jun 2014, with the prescription cost per patient ranging from 168 £26.4 to £64.5. During this time period, the highest actual cost of prescribed medications per patient for the individual practice was ~40% higher than the average 169 170 prescribing cost of $\mu = \pounds 46.1$, 95%CI = $[\pounds 45.2, \pounds 47.0]$. In addition, the average cost of prescribing per person was observed to increase by 11.3%, 95%CI = 171 172 [10.4%, 12.2%] over the period of investigation; from £45, 95%CI = [£43.2, £46.8] in the first quarter (Apr-Jun 2013) to \pounds 48.6, 95%CI = [\pounds 46.7, \pounds 50.6] in the last quarter 173 174 (Jan-Mar 2016) of the study.

Fig 2. Temporal variability in the actual cost of prescribed medications per patient. Each data point (dot): a single practice. Solid, horizontal line inside the box: median of data. Green diamond: mean. Lower and upper "hinges" of the boxplots: 1st and 3rd quartiles, respectively. Red, green, and blue lines: trend lines for maximum, average, and minimum values of prescription costs respectively.

The distribution of costs through time appeared to show a similar trend to the 180 181 prescribing rates. The moderate to strong association between prescription rates and 182 actual costs of prescribed medications was reflected in the value of the Spearman's coefficient (Fig 3 A). The rho was found to increase from 0.547 in Apr 2013 - Mar 2014 183 to 0.609 in Apr 2015 - Mar 2016. We also looked at the relationship between 184 185 prescribing rates and the actual cost per prescription. We found those two measures to be moderately correlated (Fig 3 B); the cost per prescription was shown to be lower 186 187 for practices with higher rates of prescribing.

Fig 3. The relationship between standardized number of prescriptions and: A) the actual cost of prescribed medications per patient; B) the actual cost per prescription. Our trend analysis showed that temporal changes in variability of prescribing rates and costs underwent an upward trend. Despite some temporal fluctuations in variance, the best fit line indicates that the value of σ^2 for prescribing rates increased from £0.70 in Apr-Jun 2013 to £0.77 in Jan-Mar 2016 (Fig 4). At the same time, the between-practice variability in prescribing costs increased from $\sigma^2 =$ £45.6 in Apr-Jun 2013 to $\sigma^2 =$ £53.4 in Jan-Mar 2016. The Mann–Kendall test confirmed a statistically significant upward trend in variability of GPs prescribing rates (p = 0.011) over the study duration.

Fig 4. Temporal changes in variance calculated for: A) the number of prescriptions per
patient; B) the actual prescription cost per patient for 55 investigated general practices.

199 Black line represents the best-fit trend line for rates (A) and cost (B) of prescribing.

Rural practices had a lower average number of prescriptions per patient than urban 200 201 practices at all time points (Table 1). Over the period of investigation, the mean number of prescriptions per patient for rural practices rose by ~3.3 % from 5.07, 95Cl = 202 [4.70,5.44] in Apr-Jun 2013 to 5.24, 95CI = [4.83,5.64] in Jan-Mar 2016 while urban 203 204 practices reported a ~6.7% increase in average prescribing rate from 5.14, 95CI = [4.86,5.41] in Apr-Jun 2013 to 5.48, 95CI = [5.22,5.75] in Jan-Mar 2016. In all quarterly 205 206 periods, the difference in the mean number of prescribed medications per patient between urban and rural practices was found statistically insignificant. 207

Rural practices had a higher between practice variability than urban practices at all time points (Table 1). The variance for practices designated as rural grew from $\sigma^2 =$ 0.84 in Apr-Jun 2013 to $\sigma^2 = 1.02$ in Jan-Mar 2016. This upward trend in variability was found statistically significant with p = 0.0032. Conversely, the variance for urban practices decreased from $\sigma^2 = 0.62$ in Apr-Jun 2013 to $\sigma^2 = 0.58$ in Jan-Mar 2016; however, this change was statistically insignificant (p = 0.54). At all studied time periods, F-test *p*-value showed no significant differences in variance in prescribing rates between rural and urban practices.

Table 1. Prescribing rates for rural and urban practices. T-test *p*-value refers to the significance level of differences in the mean number of prescribing rates between rural and urban practices for all considered time period. The *p*-value of F-test assesses the difference in variances in prescribing rates between rural and urban practices.

220 Practices situated in more deprived areas were found to have higher prescribing rates 221 than those located in less deprived areas although this difference was not statistically 222 significant in any of the considered quarterly periods (Table 2). The average number 223 of prescriptions per patient in less deprived areas grew by $\sim 7.5\%$ from 5.0, 95Cl = [4.38,5.61] in Apr-Jun 2013 to 5.37, 95CI = [4.76,5.98] in Jan-Mar 2016 while practices 224 225 situated in more deprived areas reported a ~4.6% increase in mean prescribing rate 226 from 5.14, 95CI = [4.92,5.38] in Apr-Jun 2013 to 5.38, 95CI = [5.13,5.62] in Jan-Mar 2016. The variability in prescribing rates for practices in less deprived areas was 227 228 substantially higher than for practices in more deprived areas and this difference in 229 variances was shown to be statistically significant for 8 guarterly periods (Apr 2013-230 Mar 2015) (Table 2).

Table 2. Prescribing rates for practices located in areas of different levels of socioeconomic deprivation. T-test *p*-value refers to the significance level of differences in mean number of prescribing rates between practices from less and more deprived areas. The *p*-value of F-test assesses the difference in variances in prescribing rates between practices from less and more deprived areas. Asterisk: Statistically significant difference (*p* < 0.05) in variability in prescribing rates.

237 **Discussion**

238 Over the period of investigation, the average between-practice variation in rates of prescribing was $\sigma^2 = 0.74$, 95%CI = [0.71, 0.77]. The prescribing rates of individual 239 practices ranged, on average, from 3.34, 95%CI = [3.26,3.42] to 8, 95%CI = 240 241 [7.86,8.14] prescriptions per patient. At the same time, the average variance of prescribing costs was $\sigma^2 = \pounds 47.6$, 95%CI = $[\pounds 44.4, \pounds 50.8]$ with actual cost of prescribed 242 medications per patient ranging, on average, from £27.2, 95%CI = [£26.1, £28.3] to 243 244 $\pounds 67.9, 95\%$ CI = [$\pounds 66.5, \pounds 69.3$]. While it may be challenging to define what represents an appropriate rate or cost of prescribing, it is certainly difficult to justify large 245 246 differences in prescribing between individual practices providing care to broadly similar 247 groups of patients within a single healthcare system.

248 It is worth highlighting that both rates and costs of prescribing observed in Northern 249 Ireland Western Health and Social Care Trust were found to be higher than the rates and costs recorded in England. In 2015, an average of 18.6 items was dispensed in 250 primary care for each patient registered with a GP practice in England [36] compared 251 252 to 21.2 items per head issued in WHSCT. In England, the cost of prescribed items was roughly £157 per patient, £5 per patient higher than in 2014. In comparison, the 253 average prescription cost per patient in WHSCT was £189.8, 95%CI = [182.9,196.7], 254 ~£7.6 higher that in 2014. Despite higher average rates of prescribing per patient, the 255 256 variation across England in the number of prescribed medications was higher than in 257 WHSCT with the prescribing rates ranging from 9.5 to 33.3 items per head in 2015. At the same time, the number of items per patient issued in WHSCT ranged from 13.7 to 258 31.7 [36]. 259

260 Since no demographic data was published alongside the GP prescribing data for

261 WHSCT, we could not estimate the effect of demographics of patient population on 262 variation in prescribing rates. Previous studies however showed that demographic characteristics of patient population did not fully explain prescribing behaviour of GPs 263 264 [21]. Among the factors related to the varying prescription activity, age of patients was most often factored into analyses of variation [37], although age alone did not account 265 266 for enough variation to develop an accurate model for predicting prescribing rates [38]. It was shown that age and gender accounted for approximately 25% of variation 267 268 [39,40] and additional demographic characteristics (e.g. mortality rates) up to 51% 269 [41].

270 Our study shows differences in both prescribing rates and between practice variation 271 in prescribing between rural and urban practices. The mean number of prescribed 272 items was higher in urban practices than in rural practices. The reasons for this are unclear and were beyond the scope of the present study. However, possible 273 274 explanations include differing patient populations in rural and urban areas, differences in practice organisation and workflow, as well as differences in characteristics of 275 general practitioners such as training, background, and age. Our results appear 276 277 consistent with previous studies. In Scotland, lower levels of prescribing of antidepressants were found for practices in rural areas while higher rates were 278 observed for urban practices [18]. In addition, lower rates of prescribing of 279 280 psychotropic drugs were reported by rural/small town practices in Denmark [19].

Our results indicate higher levels of prescribing for practices located in more deprived areas of Western Health and Social Care Trust and lower levels for practices from less deprived areas. Furthermore, the differences in variances of prescribing rates given different levels of local deprivation were found statistically significant for 8 quarterly periods. So far, several studies have demonstrated that socio-economic deprivation
can influence prescription rates for some medications, such as antidepressants and
lipid-lowering drugs. In England, the difference in the number of prescriptions between
the bottom 1% and top 1% areas by deprivation was 20% [42,43].

289 In addition, we found that the variability in prescribing rates underwent a statistically 290 significant upward trend reflecting larger deviations of prescribing rates of individual 291 practices from the mean prescribing rate. This can be related to the changes in socio-292 economic and demographic characteristics of patient populations but we also cannot 293 exclude possibility that these growing deviations may reflect growing differences in 294 guality of care leading to, in fact, avoidable increase in prescribing costs. However, 295 higher variability does not necessarily imply lower quality practice. It therefore requires 296 further inspection to determine if the patient populations associated with specific GP practices are different and have different needs. 297

298 A moderate (Apr 2013-Mar 2015) to strong (Apr 2015-Mar 2016) relationship was observed between prescription rates and actual costs of prescribing; a higher cost of 299 300 prescribed medications per patient was associated with a higher number of issued 301 items per patient. The differences in pharmaceutical costs observed for the practices 302 with similar prescription rates might be related to the type of prescribed drugs e.g. the 303 cost of one pack of Amiodarone (100mg tablets) is £2.21 whereas for a pack of 304 Allopurinol (100mg tablets), we have to pay over £35. The differences in prescription 305 costs in practices with similar prescribing rates may also be associated with the 306 medication choice i.e. a generic vs. brand name drug. There is evidence that inefficient prescribing by GPs increases NHS costs by hundreds of millions of pounds every year 307 308 [21]. Of course, there can be legitimate reasons why patients require brand name

drugs. However, our data do not allow us to examine the appropriateness of such decisions. We also found the number of items per patient to be negatively correlated with the cost per item i.e. the cost per prescription was shown to be higher for practices with lower rates of prescribing. It suggests that practices that prescribe more items per head appear to prescribe cheaper drugs.

314 We believe that the identification of outlier practices i.e. practices with higher or lower 315 prescribing rates than the calculated outlier cut-off values may act as an important 316 consideration when deciding which practices may benefit from interventions to alter 317 prescribing behaviour of GPs [44]. That is, there might be greater merit in engaging 318 with individual practices where prescribing rates appeared significantly higher or lower 319 than average. The identification of such practices could reduce the time, effort, and cost of any intervention. However, we are aware that a statistical outlier in terms of 320 prescribing rates is not equivalent to inappropriate practice and therefore, further 321 322 analysis would be required to assess if higher/lower rates than outlier cut off values 323 can be explained by characteristics of patient populations (e.g. age, ethnicity) or 324 practice features (e.g. age, training of general practitioners).

The main limitation of our study results from its design. Our analysis was conducted 325 326 to investigate the variability patterns and changes in prescribing rates and costs, but 327 due to data unavailability, we were not able to examine how the differences in patient or provider factors may affect variation in prescribing. Business Services Organisation 328 329 in Northern Ireland does not provide free and open access to data sets related to 330 demographic characteristics of patient population and practice features at the level of the GP practice. We believe that when such data becomes available, further 331 332 investigation of characteristics of practices and patient populations in Western Health and Social Care Trust may shed more light on other factors contributing to variations
 in GPs prescribing. This can help us to better address potential inappropriateness and
 inefficiency of prescribing.

336 In conclusion, our study provided information on variability patterns and temporal 337 changes in rates and cost of prescribing in Western Health and Social Care Trust. We 338 showed that practice setting and socio-economic deprivation account for some of the 339 between-practice variation in prescribing. We suggest that optimisation of prescribing 340 could be enhanced by conducting appropriate clinical interventions when other factors 341 contributing to prescribing variation are identified. These interventions could include 342 educational initiatives and feedback during which GP practices would be informed about their own frequency of prescribing relative to the mean prescribing of other 343 practices. The prescribing behaviour of GPs could also be altered by comparing their 344 past performance to clearly defined professional standards/targets. The quality 345 346 improvement initiatives including normative feedback proved to be effective in 347 decreasing variability in prescribing in the past [45].

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353 Contributorship

MB performed the analysis and interpretation of the results, and wrote the manuscript.
 MJO edited the manuscript. KWL initiated the collaborative project, guided the data

analysis and interpretation of the results, and wrote the manuscript. SA monitored
the data collection. All the authors have accepted responsibility for the entire content
of this submitted manuscript and approved submission.

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363 **Competing interests**

364 The authors have declared that no competing interests exist.

365 **References**

- Prescribing Costs in Hospitals and the Community: England 2015/16.
 Government Statistical Service. 2016. Available at:
- 368 http://www.content.digital.nhs.uk/catalogue/PUB22302/hosp-pres-eng-
- 369 <u>201516-report.pdf</u>. Accessed: 3 May 2017.
- Thompson M, Walter F. Increases in general practice workload in England. The
 Lancet 2016;387(10035):2270-2272.
- 372 3. Thompson J, Todd D. Primary Care Prescribing. Northern Ireland Assembly.
 373 2015. Available at:
- 374 https://www.degruyter.com/downloadpdf/j/cclm.2017.55.issue-5/cclm-2016-
- 375 <u>0740/cclm-2016-0740.pdf</u>. Accessed: 4 May 2017.
- 4. Prescriptions Dispensed in the Community, Statistics for England 2005-2015.
- 377 Prescribing and Medicines Team Health and Social Care Information Centre.
- 378 2016. Available at:
- 379 <u>http://content.digital.nhs.uk/catalogue/PUB20664/pres-disp-com-eng-2005-15-</u>
 380 rep.pdf. Accessed: 4 May 2017.
- 5. Duerden M, Millson D, Avery A, Smart S. The quality of GP prescribing. The
 King's Fund. 2011. Available at:
- 383 https://www.kingsfund.org.uk/sites/files/kf/field/field_document/quality-gp-
- 384 prescribing-gp-inquiry-research-paper-mar11.pdf. Accessed: 6 May 2017.
- 385 6. Alderwick H, Robertson R, Appleby J, Dunn P, Maguire D. Better value in the
- 386 NHS: The role of changes in clinical practice. The King's Fund. 2015.

- 7. Pharoah PD, Melzer D. Variation in prescribing of hypnotics, anxiolytics and
 antidepressants between 61 general practices. Br J Gen Pract.
 1995;45(400):595-9.
- Covvey JR, Johnson BF, Elliott V, Malcolm W, Mullen AB. An association
 between socioeconomic deprivation and primary care antibiotic prescribing in
 Scotland. Journal of Antimicrobial Chemotherapy. 2014;69(3):835-41.
- 393 9. Ventola CL. The Antibiotic Resistance Crisis: Part 1: Causes and
 394 Threats. Pharmacy and Therapeutics 2015;40(4):277-283.
- 395 10. Wagner AC, Hann M, Ashcroft DM. Influence of population and general practice
 396 characteristics on prescribing of minor tranquilisers in primary care. Pharmacy
 397 Practice 2010;8(3):193-200.
- 11. Hann M, Cantrill J, Baker D. Gill P. Prescribing patterns in high-need Health
 Authority populations: how does an ethnically mixed composition affect volume
 and cost? Journal of Clinical Pharmacy and Therapeutics 2004;29:537-546.
- 401 12. Backer C. Medicine statistics: GP prescribing by constituency, 2015. House of
 402 Commons Library. 2016.
- 403 13. Mackenzie IF, Buckingham K, Wankowski JM, Wilcock M. Morbidity,
 404 deprivation, and antidepressant prescribing in general practice. Br J Gen Pract.
 405 1999;49(448):884-6.
- 406 14. Wu J, Zhu S, Yao GL, Mohammed MA, Marshall T. Patient factors influencing
 407 the prescribing of lipid lowering drugs for primary prevention of cardiovascular
 408 disease in UK general practice: a national retrospective cohort study. PLoS
 409 One 2013;8(7):e67611.
- 410 15. Hull SA, Aquino P, Cotter S. Explaining variation in antidepressant prescribing
 411 rates in east London: a cross sectional study. Fam Pract 2005;22(1):37-42.

412	16. Ashworth M, Lloyd D, Smith RS, Wagner A, Rowlands G. Social deprivation
413	and statin prescribing: a cross-sectional analysis using data from the new UK
414	general practitioner 'Quality and Outcomes Framework'. J Public Health
415	2007;29(1):40-47.
416	17. Szczepura A. Access to health care for ethnic minority populations.
417	Postgraduate medical journal 2005;81(953):141-7.
418	18. Morrison J, Anderson MJ, Sutton M, Munoz-Arroyo R, McDonald S, Maxwell M,
419	et al. Factors influencing variation in prescribing of antidepressants by general
420	practices in Scotland. Br J Gen Pract 2009;59(559): e25-e31.
421	19. Holm M, Olesen F. Factors affecting prescription of psychotropic drugs in
422	general practice. Scandinavian journal of primary health care 1988;6(3):169-
423	73.
424	20. Haastrup PF, Rasmussen S, Hansen JM, Christensen RD, Søndergaard J,

- Jarbøl DE. General practice variation when initiating long-term prescribing of
 proton pump inhibitors: a nationwide cohort study. BMC family practice
 2016;17(1):57.
- 428 21.Cahir C, Fahey T, Teeling M, Teljeur C, Feely J, Bennett K. Potentially
 429 inappropriate prescribing and cost outcomes for older people: a national
 430 population study. Br J Clin Pharmacol 2010;69:543-552.
- 431 22. Ofori-Asenso R, Agyeman AA. Irrational Use of Medicines—A Summary of Key
 432 Concepts. Pharmacy. 2016 Oct 28;4(4):35.
- 433 23. Paying the Price Prescription Charges and People with Long-Term Conditions.
 434 Prescription Charges Coalition. 2013. Available at:
- 435 <u>http://www.prescriptionchargescoalition.org.uk/uploads/1/2/7/5/12754304/payi</u>
- 436 <u>ng_the_price_report.pdf</u>. Accessed: 20 April 2017.

437	24. Lash S, Harding J. Abandoned Prescriptions: A Quantitative Assessment of
438	Their Cause. Journal of Managed Care Pharmacy 1995;1(3):193-9.
439	25. Fischer SH, Field TS, Gagne SJ, Mazor KM, Preusse P, Reed G, et al. Patient
440	completion of laboratory tests to monitor medication therapy: a mixed-methods
441	study. Journal of general internal medicine 2013;28(4):513-21.
442	26. Moon JC, Flett AS, Godman BB, Grosso AM, Wierzbicki AS. Getting better
443	value from the NHS drug budget. BMJ (Online) 2010;341.
444	27. Prescribing by GP Practice. Available at:
445	http://www.hscbusiness.hscni.net/services/2471.htm. Accessed: 1 February
446	2017.
447	28. Northern Ireland Statistics and Research Agency. Review of the Statistical
448	Classification and Delineation of Settlements,
449	http://www.nisra.gov.uk/archive/geography/review-of-the-statistical-
450	classification-and-delineation-of-settlements-march-2015.pdf (2015, accessed
451	8 Nov 2017).
452	29. Northern Ireland Multiple Deprivation Measure 2010. Northern Ireland
453	Statistics and Research Agency. 2010. Available at:
454	http://www.prescriptionchargescoalition.org.uk/uploads/1/2/7/5/12754304/payi
455	ng the price report.pdf. Accessed: 8 Nov 2017.
456	30. Asimow LA, Maxwell MM. Probability and statistics with applications: A problem
457	solving text. Actex Publications. 2010.
458	31. Peck R, Olsen C, Devore JL. Introduction to statistics and data analysis.
459	Cengage Learning; 2015.
460	32. Sen A, Srivastava M. Regression analysis: theory, methods, and applications.
461	Springer Science & Business Media; 2012.

- 462 33. McLeod AI, Hipel KW, Bodo BA. Trend analysis methodology for water quality
 463 time series. Environmetrics 2010;2:169-200.
- 34. Meals DW, Spooner J, Dressing SA, Harcum JB. Statistical analysis for
 monotonic trends, Tech Notes 6. Developed for U.S. Environmental Protection
 Agency by Tetra Tech, Inc. 2011. Available at:
- 467 <u>www.bae.ncsu.edu/programs/extension/wqg/319monitoring/tech_notes.htm.</u>
- 468 Accessed: 10 May 2017.
- 35. Kvam PH, Vidakovic B. Nonparametric statistics with applications to scienceand engineering. John Wiley & Sons. 2007.
- 471 36. Baker K. Medicine statistics: GP prescribing by constituency, 2015. House of
 472 Commons Library, 2016.
- 37. Purves I, Edwards C. Comparison of prescribing unit with index including both
 age and sex in assessing general practice prescribing costs. Br Med J
 1993;306:496-8.

38. Sleator D. Towards accurate prescribing analysis in general practice: accounting for the effects of practice demography. Br J Gen Pract 1993;43(368):102-6.

- 39. Favato G, Mariani P, Mills R, Capone A, Pelagatti M, Pieri V et al. ASSET
 (Age/Sex Standardised Estimates of Treatment): a research model to improve
 the governance of prescribing funds in Italy. PLoS ONE 2007;2(7): e592.
- 482 40. Majeed A, Head S. Controversies in primary care. Setting prescribing budgets
 483 in general practice. Capitation based prescribing budgets will not work. BMJ
 484 1998;316(7133):748 –50.

485	41. Forster D, Frost C. Use of regression analysis to explain the variation in
486	prescribing rates and costs between family practitioner committees. Br J Gen
487	Pract 1991;41(343):67-71.

- 42. Hann M, Cantrill J, Baker D, Gill P. Prescribing patterns in high-need Health
 Authority populations: how does an ethnically mixed composition affect volume
 and cost? Journal of Clinical Pharmacy and Therapeutics 2004;29:537-546.
- 491 43. Wise J. Antibiotic prescribing is higher in deprived areas of England. BMJ:
 492 British Medical Journal (Online) 2005;351.
- 493 44. Hallsworth M, Chadborn T, Sallis A, Sanders M, Berry D, Greaves F, et al.
 494 Provision of social norm feedback to high prescribers of antibiotics in general
 495 practice: a pragmatic national randomised controlled trial. Lancet
 496 2016;387:1743–52.
- 497 45. Burton JH, Hoppe JA, Echternach JM, Rodgers JM, Donato M. Quality
 498 improvement initiative to decrease variability of emergency physician opioid
 499 analgesic prescribing. Western Journal of Emergency Medicine. 2016
 500 May;17(3):258.
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- 509 **Figure 1**





512 Figure 2



Figure 3







Table 1

	Apr- Jun 2013	Jul- Sep 2013	Oct- Dec 2013	Jan- Mar 2014	Apr- Jun 2014	Jul- Sep 2014	Oct- Dec 2014	Jan- Mar 2015	Apr- Jun 2015	Jul- Sep 2015	Oct- Dec 2015	Jan- Mar 2016
Rural												
Mean	5.07	5.08	5.15	5.16	5.16	5.17	5.26	5.28	5.07	5.11	5.26	5.24
Variance	0.84	0.74	0.78	0.78	0.78	0.80	0.77	0.87	0.94	0.94	1.07	1.02
Urban												
Mean	5.14	5.19	5.23	5.23	5.34	5.24	5.39	5.33	5.31	5.43	5.50	5.48
Variance	0.62	0.62	0.68	0.65	0.68	0.67	0.65	0.62	0.79	0.69	0.68	0.58
F-test p-value	0.41	0.65	0.72	0.63	0.72	0.64	0.66	0.38	0.65	0.41	0.24	0.14
T-test p-value	0.77	0.64	0.72	0.78	0.44	0.77	0.58	0.83	0.35	0.21	0.36	0.32

Table 2

	Apr- Jun 2013	Jul- Sep 2013	Oct- Dec 2013	Jan- Mar 2014	Apr- Jun 2014	Jul- Sep 2014	Oct- Dec 2014	Jan- Mar 2015	Apr- Jun 2015	Jul- Sep 2015	Oct- Dec 2015	Jan- Mar 2016
Less deprived areas												
Mean	5.00	5.01	5.12	5.13	5.13	5.11	5.26	5.23	5.22	5.22	5.42	5.37
Variance	1.27	1.25	1.39	1.37	1.37	1.41	1.31	1.33	1.49	1.41	1.41	1.25
More deprived areas												
Mean	5.14	5.18	5.22	5.22	5.30	5.24	5.36	5.34	5.20	5.31	5.39	5.38
Variance	0.55	0.51	0.53	0.51	0.54	0.52	0.53	0.56	0.68	0.65	0.70	0.65
F-test p-value	0.046*	0.031*	0.022*	0.018*	0.027*	0.018*	0.032*	0.037*	0.065	0.066	0.096	0.117
T-test p-value	0.665	0.625	0.776	0.788	0.624	0.709	0.770	0.767	0.966	0.790	0.926	0.979

S1 Table. The variance in prescribing rates for the data set with and without outlier practices.

	Apr- Jun 2013	Jul- Sep 2013	Oct- Dec 2013	Jan- Mar 2014	Apr- Jun 2014	Jul- Sep 2014	Oct- Dec 2014	Jan- Mar 2015	Apr- Jun 2015	Jul- Sep 2015	Oct- Dec 2015	Jan- Mar 2016
With outlier practices												
σ^2 (£)	0.70	0.67	0.71	0.69	0.72	0.71	0.70	0.72	0.85	0.81	0.85	0.77
Without outlier practices												
σ^2 (£)	0.52	0.39	0.59	0.43	0.45	0.44	0.49	0.50	0.60	0.58	0.59	0.56